

381 **Limitations:** The main contributions of our works are theoretical. From a theoretical point of view,
 382 the limitations of our paper are discussed in Section 5. In particular, we believe that tightening the
 383 gap between the upper and lower bounds in Nash regret for an infinite set of arms will require novel
 384 and non-trivial algorithmic ideas - we leave this as an important direction of future work.

385 **Broader Impact:** Due to the theoretical nature of this work, we do not foresee any adverse societal
 386 impact of this work.

387 A Proof of Concentration Bounds

388 **Lemma 1.** Any non-negative random variable $X \in [0, B]$ is B -sub Poisson, i.e., if mean $\mathbb{E}[X] = \mu$,
 389 then for all $\lambda \in \mathbb{R}$, we have $\mathbb{E}[e^{\lambda X}] \leq \exp(B^{-1}\mu(e^{B\lambda} - 1))$.

390 *Proof.* For random variable X we have

$$\begin{aligned} \mathbb{E}[\exp(\lambda X)] &= 1 + \sum_{i=1}^{\infty} \frac{\lambda^i \mathbb{E}[X^i]}{i!} \\ &\leq 1 + \sum_{i=1}^{\infty} \frac{\lambda^i \mathbb{E}\left[\frac{X}{B} B^i\right]}{i!} \\ &= 1 + \frac{\mathbb{E}[X]}{B} \sum_{i=1}^{\infty} \frac{\lambda^i B^i}{i!} \\ &\leq 1 + \frac{\mu}{B} (e^{\lambda B} - 1) \\ &\leq \exp\left(\frac{\mu}{B} (e^{\lambda B} - 1)\right). \end{aligned}$$

391

□

392 **Lemma 5.** Let $x_1, x_2, \dots, x_s \in \mathbb{R}^d$ be a fixed set of vectors and let r_1, r_2, \dots, r_s be independent
 393 ν -sub Poisson random variables satisfying $\mathbb{E}r_s = \langle x_s, \theta^* \rangle$ for some unknown θ^* . In that case, let
 394 matrix $\mathbf{V} = \sum_{j=1}^s x_j x_j^T$ and $\hat{\theta} = \mathbf{V}^{-1} \left(\sum_j r_j x_j \right)$ be the least squares estimator of θ^* . Consider
 395 any $z \in \mathbb{R}^d$ that satisfies $z^T \mathbf{V}^{-1} x_j \leq \gamma$ for all $j \in [s]$. Then, for any $\delta \in [0, 1]$ we have

$$\mathbb{P}\left\{ \langle z, \hat{\theta} \rangle \geq (1 + \delta) \langle z, \theta^* \rangle \right\} \leq \exp\left(-\frac{\delta^2 \langle z, \theta^* \rangle}{3\nu\gamma}\right) \quad \text{and} \quad (8)$$

$$\mathbb{P}\left\{ \langle z, \hat{\theta} \rangle \leq (1 - \delta) \langle z, \theta^* \rangle \right\} \leq \exp\left(-\frac{\delta^2 \langle z, \theta^* \rangle}{2\nu\gamma}\right) \quad (9)$$

396 *Proof.* We use \mathbf{X} to denote a matrix with arm pulls x_1, x_2, \dots, x_s stacked as rows. We use the
 397 Chernoff method to get an upper bound on the desired probabilities, as shown below

$$\begin{aligned} \mathbb{P}\left\{ \langle z, \hat{\theta} \rangle \geq (1 + \delta) \langle z, \theta^* \rangle \right\} &= \mathbb{P}\left(\exp(c \langle z, \hat{\theta} \rangle) \geq \exp(c(1 + \delta) \langle z, \theta^* \rangle)\right) \quad (\text{for some constant } c) \\ &\leq \frac{\mathbb{E}[\exp(c z^T \mathbf{V}^{-1} \mathbf{X}^T R)]}{\exp(c(1 + \delta) \langle z, \theta^* \rangle)} \\ &= \frac{\prod_{t=1}^s \mathbb{E}[\exp(c r_t \mathbf{V}^{-1} x_t)]}{\exp(c(1 + \delta) \langle z, \theta^* \rangle)} \quad (r_t \text{'s are independent}) \\ &\leq \frac{\prod_{t=1}^s \exp\left(\frac{\mathbb{E}[r_t]}{\nu} \left(e^{c \nu z^T \mathbf{V}^{-1} x_t} - 1\right)\right)}{\exp(c(1 + \delta) \langle z, \theta^* \rangle)} \quad (r_t \text{ is sub-poisson}) \\ &= \exp\left(-c \langle z, \theta^* \rangle (1 + \delta) + \sum_{t=1}^s \frac{\langle x_t, \theta^* \rangle}{\nu} \left(e^{c \nu z^T \mathbf{V}^{-1} x_t} - 1\right)\right). \end{aligned}$$

398 Substituting $c = \frac{\log(1+\delta)}{\nu\gamma}$, we get

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1+\delta)\langle z, \theta^* \rangle\right\} \leq \exp\left(-\frac{\langle z, \theta^* \rangle}{\nu\gamma}(1+\delta)\log(1+\delta) + \sum_{t=1}^s \frac{\langle x_t, \theta^* \rangle}{\nu} \left((1+\delta)^{\frac{1}{\gamma}z^T\mathbf{V}^{-1}x_t} - 1\right)\right). \quad (12)$$

399 Since $\frac{1}{\gamma}z^T\mathbf{V}^{-1}x_t \leq 1$ we have $(1+\delta)^{\frac{1}{\gamma}z^T\mathbf{V}^{-1}x_t} \leq 1 + \delta \cdot \frac{1}{\gamma}z^T\mathbf{V}^{-1}x_t$. Substituting in (12) we get

$$\begin{aligned} & \mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1+\delta)\langle z, \theta^* \rangle\right\} \\ & \leq \exp\left(-\frac{1}{\nu\gamma}\langle z, \theta^* \rangle(1+\delta)\log(1+\delta) + \sum_{t=1}^s \langle x_t, \theta^* \rangle \cdot \frac{\delta}{\nu\gamma}z^T\mathbf{V}^{-1}x_t\right) \\ & = \exp\left(-\frac{1}{\nu\gamma}\langle z, \theta^* \rangle(1+\delta)\log(1+\delta) + \frac{\delta}{\nu\gamma} \sum_{t=1}^s \theta^{*T}x_t x_t^T \mathbf{V}^{-1}z\right) \quad (\text{Rearranging terms}) \\ & = \exp\left(-\frac{1}{\nu\gamma}\langle z, \theta^* \rangle(1+\delta)\log(1+\delta) + \frac{\delta}{\nu\gamma}\langle z, \theta^* \rangle\right). \quad (\sum_{t=1}^s x_t x_t^T = \mathbf{V}) \end{aligned}$$

400 Using log inequality $\log(1+\delta) \geq \frac{2\delta}{2+\delta}$ we get

$$\begin{aligned} \mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1+\delta)\langle z, \theta^* \rangle\right\} & \leq \exp\left(-\frac{\langle z, \theta^* \rangle}{\nu\gamma} \left((1+\delta)\log(1+\delta) - \delta\right)\right) \\ & \leq \exp\left(\frac{-\delta^2\langle z, \theta^* \rangle}{(2+\delta)\nu\gamma}\right) \\ & \leq \exp\left(\frac{-\delta^2 n \langle z, \theta^* \rangle}{3\nu\gamma}\right). \quad (\text{since } \delta \in [0, 1]) \end{aligned}$$

401 We follow similar steps for the lower tail (inequality (9)) to get the following expression -

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \leq (1-\delta)\langle z, \theta^* \rangle\right\} \leq \exp\left(-\frac{1}{\nu\gamma}\langle z, \theta^* \rangle(1-\delta)\log(1-\delta) - \frac{\delta}{\nu\gamma}\langle z, \theta^* \rangle\right).$$

402 Now using inequality $(1-\delta)\log(1-\delta) \geq -\delta + \frac{\delta^2}{2}$, we get

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \leq (1-\delta)\langle z, \theta^* \rangle\right\} \leq \exp\left(\frac{-\delta^2\langle z, \theta^* \rangle}{2\nu\gamma}\right)$$

403

□

404 Combining (9) and (8) we get the following Corollary.

405 **Corollary 8.** Using notations as in Lemma 5 we have

$$\mathbb{P}\left\{|\langle z, \hat{\theta} \rangle - \langle z, \theta^* \rangle| \geq \delta\langle z, \theta^* \rangle\right\} \leq 2 \exp\left(-\frac{\delta^2\langle z, \theta^* \rangle}{3\gamma}\right) \quad (13)$$

406 The next two lemmas are variants of 5 where we bound the error in terms of α where $\alpha \geq \langle x, \theta^* \rangle$.

407 **Lemma 9.** Let $x_1, x_2, \dots, x_s \in \mathbb{R}^d$ be a fixed set of vectors and let r_1, r_2, \dots, r_s be independent
 408 ν -sub Poisson random variables satisfying $\mathbb{E}r_s = \langle x_s, \theta^* \rangle$ for some unknown θ^* . In that case, let
 409 matrix $\mathbf{V} = \sum_{j=1}^s x_j x_j^T$ and $\hat{\theta} = \mathbf{V}^{-1} \left(\sum_{j=1}^s r_j x_j\right)$ be the least squares estimator of θ^* . Consider
 410 any $z \in \mathbb{R}^d$ that satisfies $z^T \mathbf{V}^{-1} x_j \leq \gamma$ for all $j \in [s]$ and $\langle z, \theta^* \rangle \leq \alpha$. Then for any $\delta \in [0, 1]$ we
 411 have

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1+\delta)\alpha\right\} \leq e^{-\frac{\delta^2\alpha}{3\gamma\nu}} \quad (14)$$

412 *Proof.* Following similar steps as in the proof of Lemma 5

$$\begin{aligned} \mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1 + \delta)\alpha\right\} &\leq \frac{\mathbb{E}[\exp(c z^T \mathbf{V}^{-1} \mathbf{X}^T R)]}{\exp(c(1 + \delta)\alpha)} \\ &\leq \exp\left(-c\alpha(1 + \delta) + \sum_{t=1}^s \frac{\langle X_t, \theta^* \rangle}{\nu} \left(e^{c\nu z^T \mathbf{V}^{-1} x_t} - 1\right)\right) \end{aligned}$$

(r_t are sub-poisson and conditionally independent)

413 Now, substituting $c = \frac{1}{\nu\gamma} \log(1 + \delta)$ and using $(1 + \delta)^{\frac{1}{\gamma} z^T \mathbf{V}^{-1} x_t} \leq 1 + \delta \cdot \frac{1}{\gamma} z^T \mathbf{V}^{-1} x_t$ we have

$$\begin{aligned} \mathbb{P}\left\{\langle z, \hat{\theta} \rangle \geq (1 + \delta)\alpha\right\} &\leq \exp\left(-\frac{1}{\nu\gamma}\alpha(1 + \delta)\log(1 + \delta) + \sum_{t=1}^s \frac{\langle X_t, \theta^* \rangle}{\nu} \left((1 + \delta)^{\frac{1}{\gamma} z^T \mathbf{V}^{-1} x_t} - 1\right)\right) \\ &\leq \exp\left(-\frac{1}{\nu\gamma}\alpha(1 + \delta)\log(1 + \delta) + \frac{\delta}{\nu\gamma} \sum_{t=1}^s \theta^{*T} x_t x_t^T \mathbf{V}^{-1} Z\right) \\ &= \exp\left(-\frac{1}{\nu\gamma}\alpha(1 + \delta)\log(1 + \delta) + \frac{\delta}{\nu\gamma} \langle z, \theta^* \rangle\right) \\ &\leq \exp\left(-\frac{1}{\nu\gamma}\alpha(1 + \delta)\log(1 + \delta) + \frac{\delta}{\nu\gamma}\alpha\right) \quad (\alpha \geq \langle z, \theta^* \rangle) \\ &\leq \exp\left(\frac{-\delta^2\alpha}{(2 + \delta)\nu\gamma}\right) \quad (\text{Using } \log(1 + \delta) \geq \frac{2\delta}{2 + \delta}) \end{aligned}$$

414 Since $\delta \in [0, 1]$, we have the desired result. \square

415 **Lemma 10.** Let $x_1, x_2, \dots, x_s \in \mathbb{R}^d$ be a fixed set of vectors and let r_1, r_2, \dots, r_s be independent
416 ν -sub Poisson random variables satisfying $\mathbb{E}r_s = \langle x_s, \theta^* \rangle$ for some unknown θ^* . In that case, let
417 matrix $\mathbf{V} = \sum_{j=1}^s x_j x_j^T$ and $\hat{\theta} = \mathbf{V}^{-1} \left(\sum_j r_j x_j\right)$ be the least squares estimator of θ^* . Consider
418 any $z \in \mathbb{R}^d$ that satisfies $z^T \mathbf{V}^{-1} x_j \leq \gamma$ for all $j \in [s]$ and $\langle z, \theta^* \rangle \leq \alpha$. Then for any $\delta \in [0, 1]$ we
419 have

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \leq \langle z, \theta^* \rangle - \delta\alpha\right\} \leq \exp\left(-\frac{\delta^2\alpha}{2\gamma\nu}\right) \quad (15)$$

420 *Proof.* Using the steps as in the previous lemmas, we arrive at

$$\mathbb{P}\left\{\langle z, \hat{\theta} \rangle \leq \langle z, \theta^* \rangle - \delta\alpha\right\} \leq \exp\left(-\frac{\langle z, \theta^* \rangle}{\nu\gamma} (\log(1 - \delta) + \delta) + \frac{\alpha}{\nu\gamma} \delta \log(1 - \delta)\right)$$

421 Note that since $\log(1 - \delta) + \delta$ is negative, we can upper bound the above expression by replacing
422 $\langle z, \theta^* \rangle$ with α .

$$\begin{aligned} \mathbb{P}\left\{\langle z, \hat{\theta} \rangle \leq \langle z, \theta^* \rangle - \delta\alpha\right\} &\leq \exp\left(-\frac{\alpha}{\nu\gamma} (\log(1 - \delta) + \delta - \delta \log(1 - \delta))\right) \\ &\leq \exp\left(-\frac{\delta^2\alpha}{2\nu\gamma}\right) \quad (\text{since } (1 - \delta) \log(1 - \delta) \geq -\delta + \frac{\delta^2}{2}) \end{aligned}$$

423 \square

424 B Regret Analysis of Algorithm 2

425 Let us define events E_1 and E_2 for each phase of the algorithm and show that they hold with high
426 probability. We will use the events in the missing proofs from Section 3.3.

E_1 At the end of Part I, let $\hat{\theta}$ be the unbiased estimator of θ^* . All arms $x \in \mathcal{X}$ with $\langle x, \theta^* \rangle <$
 $10\sqrt{\frac{d \log(\Gamma|\mathcal{X}|)}{\Gamma}}$ satisfy

$$\langle x, \hat{\theta} \rangle \leq 20\sqrt{\frac{d \log(\Gamma|\mathcal{X}|)}{\Gamma}}$$

427 and arms x with $\langle x, \theta^* \rangle \geq 10\sqrt{\frac{d \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}}$ satisfy

$$\begin{aligned} |\langle x, \theta^* \rangle - \langle x, \hat{\theta} \rangle| &\leq 3\sqrt{\frac{d\langle x, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}} \\ \frac{1}{2}\langle x, \theta^* \rangle &\leq \langle x, \hat{\theta} \rangle \leq \frac{4}{3}\langle x, \theta^* \rangle. \end{aligned}$$

428 E_2 : Let $\tilde{\mathcal{X}}$ denote the candidate set at the start of a phase in Part II, and \mathbb{T}' be as defined in Algorithm

429 2. For all phases and for all $z \in \tilde{\mathcal{X}}$ such that $\langle z, \theta^* \rangle \geq 10\frac{\sqrt{d \log(\mathbb{T}|\mathcal{X}|)}}{\sqrt{\mathbb{T}}}$, the estimator $\hat{\theta}$
430 (calculated at the end of the phase) satisfies

$$\begin{aligned} |\langle z, \theta^* \rangle - \langle z, \hat{\theta} \rangle| &\leq 3\sqrt{\frac{d\langle z, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'}} \\ \frac{1}{2}\langle z, \theta^* \rangle &\leq \langle z, \hat{\theta} \rangle \leq \frac{4}{3}\langle z, \theta^* \rangle. \end{aligned}$$

431 **Lemma 11** (Chernoff Bound). Let Z_1, \dots, Z_n be independent Bernoulli random variables. Consider
432 the sum $S = \sum_{r=1}^n Z_r$ and let $\nu = \mathbb{E}[S]$ be its expected value. Then, for any $\varepsilon \in [0, 1]$, we have

$$\mathbb{P}\{S \leq (1 - \varepsilon)\nu\} \leq \exp\left(-\frac{\nu\varepsilon^2}{2}\right)$$

433 **Lemma 12**. During Part I, arms from D -optimal design are added to S at least $\tilde{\mathbb{T}}/3$ times with
434 probability greater than $1 - \frac{1}{\mathbb{T}}$

435 *Proof.* We use Lemma 11 with Z_i as indicator random variables, that take value one when an arm for
436 \mathcal{A} (the support of λ in the optimal design) is chosen. Taking $\varepsilon = \frac{1}{3}$ and $\nu = \frac{\tilde{\mathbb{T}}}{2}$ we get the required
437 probability bound. \square

Lemma 13. Using the notation in Algorithm 1 for $z \in \mathcal{X}$ we have

$$z^T \mathbf{V}^{-1} X_t \leq \frac{3d}{\mathbb{T}}$$

438 *Proof.* Let $\mathbf{U}(\lambda)$ and λ be the optimal design matrix (as defined in 4) and the solution to the
439 D -optimal design problem in Algorithm ?? i.e. if λ is the solution of the objective function in
440 equation 5 then, $\mathbf{U}(\lambda) = \sum_{x \in \mathcal{X}} \lambda_x x x^T$. Clearly, from Lemma 2 we must have that for any $z \in \mathcal{X}$,
441 $\|z\|_{\mathbf{U}(\lambda)^{-1}} \leq d$. By construction of the sequence S in Step 1 (Subroutine GenerateArmSequence),
442 we have $\mathbf{V} \succ \frac{\tilde{\mathbb{T}}}{3} \mathbf{U}(\lambda)$. Hence

$$\begin{aligned} z^T \mathbf{V}^{-1} X_t &\leq \|z\|_{\mathbf{V}^{-1}} \|\mathbf{V}^{-1} X_t\|_{\mathbf{V}} && \text{(By Hölder's inequality)} \\ &= \|z\|_{\mathbf{V}^{-1}} \|X_t\|_{\mathbf{V}^{-1}} \\ &\leq \|z\|_{\left(\frac{\tilde{\mathbb{T}}}{3} \mathbf{U}(\lambda)\right)^{-1}} \|X_t\|_{\left(\frac{\tilde{\mathbb{T}}}{3} \mathbf{U}(\lambda)\right)^{-1}} && \text{(since } \mathbf{V} \succ \frac{\tilde{\mathbb{T}}}{3} \mathbf{U}(\lambda)\text{)} \\ &= \sqrt{\frac{3}{\tilde{\mathbb{T}}}} \|z\|_{\mathbf{U}(\lambda)^{-1}} \sqrt{\frac{3}{\tilde{\mathbb{T}}}} \|X_t\|_{\mathbf{U}(\lambda)^{-1}} \\ &\leq \sqrt{\frac{3d}{\tilde{\mathbb{T}}}} \sqrt{\frac{3d}{\tilde{\mathbb{T}}}} && \text{(by Lemma } \span style="border: 1px solid red; padding: 0 2px;">2\text{)} \\ &= \frac{3d}{\tilde{\mathbb{T}}}. \end{aligned}$$

443 \square

444 **Lemma 14**. Let $\hat{\theta}$ be the estimate computed at the end of Part I of Algorithm 2 Following holds with
445 probability greater than $1 - \frac{4}{\mathbb{T}}$.

446 • All arms $x \in \mathcal{X}$ with $\langle x, \theta^* \rangle \leq 10\sqrt{d\nu\bar{T}^{-1}\log(\bar{T}|\mathcal{X}|)}$ satisfy

$$\langle x, \hat{\theta} \rangle \leq 20\sqrt{d\nu\bar{T}^{-1}\log(\bar{T}|\mathcal{X}|)}. \quad (16)$$

447 • All arms $x \in \mathcal{X}$ with $\langle x, \theta^* \rangle \geq 10\sqrt{d\nu\bar{T}^{-1}\log(\bar{T}|\mathcal{X}|)}$ satisfy

$$|\langle x, \theta^* \rangle - \langle x, \hat{\theta} \rangle| \leq 3\sqrt{\frac{d\langle x, \theta^* \rangle \log(\bar{T}|\mathcal{X}|)}{\bar{T}'}} \quad \text{and} \quad (17)$$

$$\frac{1}{2}\langle x, \theta^* \rangle \leq \langle x, \hat{\theta} \rangle \leq \frac{4}{3}\langle x, \theta^* \rangle. \quad (18)$$

448 *Proof.* First, consider the set \mathcal{X}_{low} . We use Lemma 9 for the proof. We set $\gamma = \frac{3d}{\bar{T}}$ (from Lemma 13),
 449 $\alpha = 10\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\bar{T}}}$ and $\delta = 1$,

$$\begin{aligned} \mathbb{P} \left\{ \langle x, \hat{\theta} \rangle \leq 20\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\bar{T}}} \right\} &\leq e^{-\frac{\delta^2\alpha}{3\gamma\nu}} \\ &\leq \exp \left(-\frac{3\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\bar{T}}} 3\sqrt{\bar{T}d\nu\log(\bar{T}|\mathcal{X}|)}}{3\nu d} \right) \\ &\leq \frac{1}{\bar{T}|\mathcal{X}|}. \end{aligned}$$

450 Next, we make use of Lemma 5 for (17). We set $\gamma = \frac{3d}{\bar{T}}$ and $\delta = 3\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\langle x, \theta^* \rangle \bar{T}}}$. Note that since
 451 $\langle x, \theta^* \rangle \geq 10\sqrt{d\nu\bar{T}^{-1}\log(\bar{T}|\mathcal{X}|)}$ and $\bar{T} = 3\sqrt{\bar{T}d\nu\log(\bar{T}|\mathcal{X}|)}$, δ always lies in $[0, 1]$. Hence we
 452 can apply Lemma 5 as follows

$$\begin{aligned} \mathbb{P} \left\{ |\langle X, \theta^* \rangle - \langle X, \hat{\theta} \rangle| \geq 3\sqrt{\frac{\nu d \langle x, \theta^* \rangle \log(\bar{T}|\mathcal{X}|)}{\bar{T}}} \right\} \\ \leq 2 \exp \left(-\frac{\frac{9d\nu\log(\bar{T}|\mathcal{X}|)}{\langle x, \theta^* \rangle \bar{T}} \cdot \langle x, \theta^* \rangle}{3\nu \frac{3d}{\bar{T}}} \right) \\ = \frac{2}{\bar{T}|\mathcal{X}|}. \end{aligned}$$

453 Next, we prove (18). The upper tail is obtained by setting $\gamma = \frac{3d}{\bar{T}}$, $\delta = \frac{1}{3}$ in expression (8) of Lemma
 454 5, we get

$$\begin{aligned} \mathbb{P} \left\{ \langle X, \hat{\theta} \rangle \geq \frac{4}{3}\langle x, \theta^* \rangle \right\} &\leq \exp \left(-\frac{3\sqrt{\bar{T}\nu d \log(\bar{T}|\mathcal{X}|)} \cdot 10\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\bar{T}}}}{27\nu d} \right) \\ &\quad \text{(Since } \langle x, \theta^* \rangle \geq 10\sqrt{\frac{d\nu\log(\bar{T}|\mathcal{X}|)}{\bar{T}}}) \\ &\leq \frac{1}{\bar{T}|\mathcal{X}|}. \end{aligned}$$

455 Similarly substituting $\delta = 1/2$ in expression (9) of Lemma 5 we get

$$\mathbb{P} \left\{ \langle X, \hat{\theta} \rangle \leq \frac{1}{2}\langle x, \theta^* \rangle \right\} \leq \frac{1}{\bar{T}|\mathcal{X}|}.$$

456 Union bound over all arms in \mathcal{X} gives us the required probability bound. \square

457 Next, we look at Part II of Algorithm 2 and show that the event E_2 holds with high probability. Note
 458 that since we find a sparse λ (with support size almost $\frac{d(d+1)}{2}$) in every phase, the phase length is
 459 upper bounded as $\bar{T}' + \frac{d(d+1)}{2}$.

460 **Lemma 15.** Using the notation in Algorithm 2 For all arms $x \in \tilde{\mathcal{X}}$ with $\langle x, \theta^* \rangle \geq 10 \frac{\sqrt{d\nu \log(\mathbb{T}|\mathcal{X}|)}}{\sqrt{\mathbb{T}}}$,
 461 the following holds (for every phase) with probability greater than $1 - \frac{3 \log \mathbb{T}}{\mathbb{T}}$

$$|\langle x, \theta^* \rangle - \langle x, \hat{\theta} \rangle| \leq 3 \sqrt{\frac{d\nu \langle x, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'}} \quad (19)$$

$$\frac{1}{2} \langle x, \theta^* \rangle \leq \langle x, \hat{\theta} \rangle \leq \frac{4}{3} \langle x, \theta^* \rangle \quad (20)$$

462 *Proof.* The proof follows the same structure as the proof of Lemma 14. Consider any Phase in Part
 463 II and let $\mathbf{U}(\lambda)$ be the optimal design matrix obtained after solving the D-optimal design problem at
 464 the start of the phase. Since each arm a in the support of λ (denoted by \mathcal{A}) is pulled at least $\lceil \lambda_a \mathbb{T}' \rceil$
 465 times, we have $\mathbf{V} \succ \frac{\mathbb{T}'}{3} \mathbf{U}(\lambda)$. Thus by Theorem 2, for $x \in \mathcal{A}$ and all $z \in \tilde{\mathcal{X}}$ we have

$$\begin{aligned} z^T \mathbf{V}^{-1} x &\leq \|z\|_{\mathbf{V}^{-1}} \|\mathbf{V}^{-1} x\|_{\mathbf{V}} && \text{(By Hölder's inequality)} \\ &\leq \|z\|_{\mathbf{V}^{-1}} \|x\|_{\mathbf{V}^{-1}} && (21) \end{aligned}$$

$$\leq \sqrt{\frac{d}{\mathbb{T}'}} \sqrt{\frac{d}{\mathbb{T}'}} = \frac{d}{\mathbb{T}'} \quad (22)$$

466 Now we use Lemma 5 with $\delta = 3 \sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\langle x, \theta^* \rangle \mathbb{T}'}}$ and $\gamma = \frac{d}{\mathbb{T}'}$. Note that given the lower bound on
 467 $\langle x, \theta^* \rangle$ and $\mathbb{T}' \geq 2 \sqrt{\mathbb{T} d \nu \log(\mathbb{T}|\mathcal{X}|)}$ in every phase, δ always lies in $[0, 1]$. Substituting in Lemma 5,
 468 we get

$$\begin{aligned} \mathbb{P} \left\{ |\langle X, \theta^* \rangle - \langle X, \hat{\theta} \rangle| \geq 3 \sqrt{\frac{d\nu \langle x, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'}} \right\} &\leq 2 \exp \left(- \frac{\frac{9d \log(\mathbb{T}|\mathcal{X}|)}{\langle x, \theta^* \rangle \mathbb{T}'} \cdot \langle x, \theta^* \rangle}{3 \frac{d}{\mathbb{T}'}} \right) \\ &\leq \frac{2}{(\mathbb{T}|\mathcal{X}|)^3} \end{aligned}$$

469 Similar to the proof of Lemma 14, we use Lemma 5 with $\delta = \frac{1}{3}$ and $\delta = \frac{1}{2}$ to bound the upper and
 470 lower tails of (20) respectively. Furthermore, a union bound across arms in $\tilde{\mathcal{X}}$ and all – at most $\log \mathbb{T}$
 471 – phases gives us the desired probability bound of $1 - \frac{3 \log \mathbb{T}}{\mathbb{T}}$. \square

Corollary 16.

$$\mathbb{P} \{E_1 \cap E_2\} \geq 1 - \frac{4 \log \mathbb{T}}{\mathbb{T}}.$$

472 *Proof.* From Lemma 14 we have $\mathbb{P} \{E_1\} \geq 1 - \frac{4}{\mathbb{T}}$. Furthermore from Lemma 15 we have $\mathbb{P} \{E_2\} \geq$
 473 $1 - \frac{3 \log \mathbb{T}}{\mathbb{T}}$. Taking union bound over the complements of the two events proves the corollary. \square

474 **Lemma 17.** Consider an instance with $\langle x^*, \theta^* \rangle \geq 192 \sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}}$. If E_1 holds, then any arm
 475 with mean $\langle x, \theta^* \rangle \leq 10 \sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}}$ is eliminated after Part I of Algorithm 2

476 *Proof.* From Lemma 14 for any arm with $\langle x, \theta^* \rangle \leq 10\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}}$ we have,

$$\begin{aligned}
\text{UNCB} \left(x, \hat{\theta}, \tilde{\mathbb{T}}/3 \right) &= \langle x, \hat{\theta} \rangle + 6\sqrt{\frac{3\langle x, \hat{\theta} \rangle d\nu \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} \\
&\leq 20\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} + 6\sqrt{\frac{3\langle x, \hat{\theta} \rangle d\nu \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} \\
&\leq 20\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} + 6\sqrt{\frac{3 \cdot 20\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} d\nu \log(\mathbb{T}|\mathcal{X}|)}{3\sqrt{\mathbb{T}d\nu \log(\mathbb{T}|\mathcal{X}|)}}} \\
&\hspace{15em} \text{(via Lemma 14 } \langle x, \hat{\theta} \rangle \leq 20\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} \text{)} \\
&\leq 47\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}}. \tag{23}
\end{aligned}$$

477 For the optimal arm x^* we have

$$\begin{aligned}
\langle x^*, \hat{\theta} \rangle &\leq \langle x^*, \theta^* \rangle + 3\sqrt{\frac{d\nu \langle x^*, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} = \langle x^*, \theta^* \rangle \left(1 + 3\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\langle x^*, \theta^* \rangle 3\sqrt{\mathbb{T}d\nu \log(\mathbb{T}|\mathcal{X}|)}}} \right) \\
&\hspace{15em} \text{(Substituting the value of } \tilde{\mathbb{T}} \text{)} \\
&\leq \langle x^*, \theta^* \rangle \left(1 + 3\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{192\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} 3\sqrt{\mathbb{T}d\nu \log(\mathbb{T}|\mathcal{X}|)}}} \right) \\
&= \frac{17}{16} \langle x^*, \theta^* \rangle. \tag{24}
\end{aligned}$$

478 This gives us a lower bound on the LNCB of x^*

$$\begin{aligned}
\text{LNCB} \left(x^*, \hat{\theta}, \tilde{\mathbb{T}}/3 \right) &= \langle x^*, \hat{\theta} \rangle - 6\sqrt{\frac{3\langle x^*, \hat{\theta} \rangle d\nu \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} \\
&\geq \langle x^*, \theta^* \rangle - 3\sqrt{\frac{d\nu \langle x^*, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} - 6\sqrt{\frac{3\langle x^*, \hat{\theta} \rangle d\nu \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} \\
&\hspace{15em} \text{(via Lemma 14)} \\
&\geq \langle x^*, \theta^* \rangle - \left(3 + 6\sqrt{\frac{51}{16}} \right) \sqrt{\frac{d\nu \langle x^*, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\tilde{\mathbb{T}}}} \\
&\hspace{15em} \text{(since } \langle x^*, \hat{\theta} \rangle \leq \frac{17}{16} \langle x^*, \theta^* \rangle \text{)} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - 14\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \tilde{\mathbb{T}}}} \right) \\
&\geq \langle x^*, \theta^* \rangle \left(1 - 14\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{192\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{T}} 3\sqrt{\mathbb{T}d\nu \log(\mathbb{T}|\mathcal{X}|)}}} \right) \\
&\geq \frac{5}{12} \langle x^*, \theta^* \rangle \\
&\geq 80\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}}}. \tag{25}
\end{aligned}$$

479 From (25) and (23) we have

$$\text{UNCB} \left(x, \hat{\theta}, \tilde{\mathbb{T}}/3 \right) \leq \text{LNCB} \left(x^*, \hat{\theta}, \tilde{\mathbb{T}}/3 \right). \tag{26}$$

480 \square

481 **Lemma 6.** *The optimal arm x^* always exists in the surviving set $\tilde{\mathcal{X}}$ in Part I and in every phase in*
 482 *Part II of Algorithm 2 with probability at least $1 - O(T^{-1} \log T)$.*

483 *Proof.* Let us assume that events E_1 and E_2 hold. For any arm x in \mathcal{X} with $\langle x, \theta^* \rangle \geq$
 484 $10\sqrt{\frac{d\nu \log(T|\mathcal{X}|)}{T}}$, we have

$$\begin{aligned} \text{LNCB}(x, \hat{\theta}, T') &= \langle x, \hat{\theta} \rangle - 6\sqrt{\frac{\langle x, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{T'}} \\ &\leq \langle x, \theta^* \rangle + 3\sqrt{\frac{d\nu \langle x, \theta^* \rangle \log(T|\mathcal{X}|)}{T'}} - 6\sqrt{\frac{\langle x, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{T'}} \\ &\leq \langle x, \theta^* \rangle - \left(\frac{6}{\sqrt{2}} - 3\right) \sqrt{\frac{d\nu \langle x, \theta^* \rangle \log(T|\mathcal{X}|)}{T'}} \\ &\leq \langle x, \theta^* \rangle. \end{aligned}$$

485 Similarly, we have

$$\begin{aligned} \text{UNCB}(x^*, \hat{\theta}, T') &= \langle x^*, \hat{\theta} \rangle + 6\sqrt{\frac{\langle x^*, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{T'}} \\ &\geq \langle x^*, \theta^* \rangle - 3\sqrt{\frac{d\nu \langle x^*, \theta^* \rangle \log(T|\mathcal{X}|)}{\tilde{T}}} + 6\sqrt{\frac{\langle x^*, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{T'}} \\ &\geq \langle x^*, \theta^* \rangle + \left(\frac{6}{\sqrt{2}} - 3\right) \sqrt{\frac{d\nu \langle x^*, \theta^* \rangle \log(T|\mathcal{X}|)}{\tilde{T}}} \\ &\geq \langle x^*, \theta^* \rangle. \end{aligned}$$

486 Since $\langle x^*, \theta^* \rangle \geq \langle x, \theta^* \rangle \forall x \in \mathcal{X}$, we have $\text{UNCB}(x^*, \hat{\theta}, T') \geq \text{LNCB}(x, \hat{\theta}, T') \forall \mathcal{X}$. From
 487 Corollary 16, we have that the events E_1 and E_2 hold with probability greater than $1 - \frac{4\log T}{T}$. Hence,
 488 the lemma stands proven. \square

489 **Lemma 7.** *Consider any phase ℓ in Part II of Algorithm 2 and let $\tilde{\mathcal{X}}$ be the surviving set of arms at*
 490 *the beginning of that phase. Then, with $\tilde{T} = \sqrt{d\nu T \log(T|\mathcal{X}|)}$, we have*

$$\Pr \left\{ \langle x, \theta^* \rangle \geq \langle x^*, \theta^* \rangle - 25\sqrt{\frac{3d\nu \langle x^*, \theta^* \rangle \log(T|\mathcal{X}|)}{2^\ell \cdot \tilde{T}}} \text{ for all } x \in \tilde{\mathcal{X}} \right\} \leq 4T^{-1} \log T \quad (10)$$

491 Here, ν is the sub-Poisson parameter of the stochastic rewards.

492 *Proof.* Let us assume that events E_1 and E_2 hold. From the second phase onwards, if an arm is
 493 pulled in a phase with phase length parameter T' , then it was not eliminated in the previous phase
 494 with phase length parameter $\frac{T'}{2}$. Additionally, since the best arm is always present in the surviving
 495 arm set $\tilde{\mathcal{X}}$ (via Lemma 6), we have $\text{UNCB}(x, \hat{\theta}, T'/2) \geq \text{LNCB}(x^*, \hat{\theta}, T'/2)$. That is

$$\langle x, \hat{\theta} \rangle + 6\sqrt{\frac{\langle x, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{\frac{T'}{2}}} \geq \langle x^*, \hat{\theta} \rangle - 6\sqrt{\frac{\langle x^*, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{\frac{T'}{2}}}.$$

496 Rearranging terms, we get

$$\begin{aligned} \langle x, \hat{\theta} \rangle &\geq \langle x^*, \hat{\theta} \rangle - 6\sqrt{\frac{\langle x^*, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{\frac{T'}{2}}} - 6\sqrt{\frac{\langle x, \hat{\theta} \rangle d\nu \log(T|\mathcal{X}|)}{\frac{T'}{2}}} \\ &\geq \langle x^*, \hat{\theta} \rangle - 6\sqrt{\frac{4\langle x^*, \theta^* \rangle d\nu \log(T|\mathcal{X}|)}{T'}} - 6\sqrt{\frac{4\langle x, \theta^* \rangle d\nu \log(T|\mathcal{X}|)}{T'}} \\ &\quad \text{(via Lemma 7 all surviving arms satisfy } \langle x, \hat{\theta} \rangle \leq \frac{4}{3}\langle x, \theta^* \rangle) \\ &\geq \langle x^*, \hat{\theta} \rangle - 20\sqrt{\frac{\langle x^*, \theta^* \rangle d\nu \log(T|\mathcal{X}|)}{T'}}. \end{aligned}$$

497 Now using the additive confidence intervals we have,

$$\begin{aligned} \langle x, \theta^* \rangle &\geq \langle x^*, \theta^* \rangle - 20\sqrt{\frac{\langle x^*, \theta^* \rangle d \nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'}} - 3\sqrt{\frac{\langle x^*, \theta^* \rangle d \nu \log(\mathbb{T}|\mathcal{X}|)}{\frac{\mathbb{T}'}{2}}} \\ &\geq \langle x^*, \theta^* \rangle - 25\sqrt{\frac{\langle x^*, \theta^* \rangle d \nu \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'}}. \end{aligned}$$

498 Substituting $\mathbb{T}' = 2^l \tilde{\mathbb{T}}/3$ in the above inequality proves the Lemma. From Corollary [16](#), we have
 499 that the events E_1 and E_2 hold with probability greater than $1 - \frac{4\log \mathbb{T}}{\mathbb{T}}$. Hence, the lemma stands
 500 proven. \square

501 **Theorem 1.** Consider the stochastic linear bandits problem over a horizon of \mathbb{T} rounds such that at
 502 every round $t \in [\mathbb{T}]$, an arm $X_t \in \mathcal{X} \subset \mathbb{R}^d$ is selected and the corresponding reward r_t is obtained
 503 satisfying equation [\(2\)](#). In the setting when \mathcal{X} is finite, Algorithm [2](#) achieves a Nash regret of

$$\text{NR}_{\mathbb{T}} = O\left(\sqrt{\frac{d\nu\langle x^*, \theta^* \rangle}{\mathbb{T}} \log(\mathbb{T}|\mathcal{X}|)}\right).$$

504 *Proof.* WLOG we assume that $\langle x^*, \theta^* \rangle \geq 192\sqrt{\frac{d\nu}{\mathbb{T}}} \log(\mathbb{T}|\mathcal{X}|)$, otherwise the Nash Regret bound is
 505 trivially true. During Part I of Algorithm [2](#) the product of expected rewards, conditioned on the event
 506 $E_1 \cap E_2$, satisfies

$$\begin{aligned} \prod_{t=1}^{\tilde{\mathbb{T}}} \mathbb{E}[\langle X_t, \theta^* \rangle \mid E_1 \cap E_2]^{\frac{1}{\tilde{\mathbb{T}}}} &\geq \left(\frac{\langle x^*, \theta^* \rangle}{2(d+1)}\right)^{\frac{\tilde{\mathbb{T}}}{\tilde{\mathbb{T}}}} && \text{(From Lemma [4](#))} \\ &= \langle x^*, \theta^* \rangle^{\frac{\tilde{\mathbb{T}}}{\tilde{\mathbb{T}}}} \left(1 - \frac{1}{2}\right)^{\frac{\log(2(d+1))\tilde{\mathbb{T}}}{\tilde{\mathbb{T}}}} \\ &\geq \langle x^*, \theta^* \rangle^{\frac{\tilde{\mathbb{T}}}{\tilde{\mathbb{T}}}} \left(1 - \frac{\log(2(d+1))\tilde{\mathbb{T}}}{\mathbb{T}}\right). \end{aligned}$$

507 For Part II, we use Lemma [7](#). Let set \mathcal{E}_i denote all t that belong to i^{th} phase and let \mathbb{T}'_i be the phase
 508 length parameter in that phase. Since each arm x in \mathcal{A} (the support of D-optimal design) is pulled
 509 $\lceil \lambda_x \mathbb{T}'_i \rceil$ times, we have $|\mathcal{E}_i| \leq \mathbb{T}'_i + \frac{d(d+1)}{2}$. Since the phase length parameter doubles after phase,
 510 the algorithm would have at most $\log \mathbb{T}$ phases. Hence we have

$$\begin{aligned} \prod_{t=\tilde{\mathbb{T}}+1}^{\mathbb{T}} \mathbb{E}[\langle X_t, \theta^* \rangle \mid E_1 \cap E_2]^{\frac{1}{\mathbb{T}}} &= \prod_{\mathcal{E}_j} \prod_{t \in \mathcal{E}_j} \mathbb{E}[\langle X_t, \theta^* \rangle \mid E_1 \cap E_2]^{\frac{1}{\mathbb{T}}} \\ &= \prod_{\mathcal{E}_j} \prod_{t \in \mathcal{E}_j} \mathbb{E}[\langle X_t, \theta^* \rangle \mid E_1 \cap E_2]^{\frac{1}{\mathbb{T}}} \\ &\geq \prod_{\mathcal{E}_j} \left(\langle x^*, \theta^* \rangle - 25\sqrt{\frac{d\nu\langle x^*, \theta^* \rangle \log(\mathbb{T}|\mathcal{X}|)}{\mathbb{T}'_j}} \right)^{\frac{|\mathcal{E}_j|}{\mathbb{T}}} \\ &\geq \langle x^*, \theta^* \rangle^{\frac{\mathbb{T}-\tilde{\mathbb{T}}}{\mathbb{T}}} \prod_{i=1}^{\log \mathbb{T}} \left(1 - 25\sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right)^{\frac{|\mathcal{E}_j|}{\mathbb{T}}} \\ &\geq \langle x^*, \theta^* \rangle^{\frac{\mathbb{T}-\tilde{\mathbb{T}}}{\mathbb{T}}} \prod_{i=1}^{\log \mathbb{T}} \left(1 - 50\frac{|\mathcal{E}_j|}{\mathbb{T}} \sqrt{\frac{d\nu \log(\mathbb{T}|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right). \end{aligned}$$

511 The last inequality is due to the fact that $(1-x)^r \geq (1-2rx)$ where $r \in [0, 1]$ and $x \in [0, 1/2]$.
 512 Note that the term $\sqrt{\frac{d\log(\mathbb{T}|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \leq 1/2$ for $\langle x^*, \theta^* \rangle \geq 192\sqrt{\frac{d}{\mathbb{T}}} \log(\mathbb{T}|\mathcal{X}|)$, $\mathbb{T}' \geq 2\sqrt{\mathbb{T}d \log \mathbb{T}|\mathcal{X}|}$
 513 and $\mathbb{T} \geq e^4$. We now further simplify the expression as shown below

$$\begin{aligned}
\prod_{j=1}^{\log T} \left(1 - 50 \frac{|\mathcal{E}_j|}{T} \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \Gamma'_j}} \right) &\geq \prod_{j=1}^{\log T} \left(1 - 50 \frac{\Gamma'_j + \frac{d(d+1)}{2}}{T} \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{\langle x^*, \theta^* \rangle \Gamma'_j}} \right) \\
&\geq \prod_{j=1}^{\log T} \left(1 - 75 \frac{\sqrt{\Gamma'_j}}{T} \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{\langle x^*, \theta^* \rangle}} \right) \\
&\hspace{15em} \text{(assuming } \Gamma'_j \geq d(d+1) \text{)} \\
&\geq 1 - \frac{75}{T} \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{\langle x^*, \theta^* \rangle}} \left(\sum_{j=1}^{\log T} \sqrt{\Gamma'_j} \right) \\
&\hspace{15em} \text{(since } (1-a)(1-b) \geq 1-a-b \text{ } \forall a, b \geq 0 \text{)} \\
&\geq 1 - \frac{75}{T} \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{\langle x^*, \theta^* \rangle}} \left(\sqrt{T \log T} \right) \\
&\hspace{15em} \text{(using Cauchy Schwarz)} \\
&\geq 1 - 75 \sqrt{\frac{d \nu}{T \langle x^*, \theta^* \rangle}} \log(T|\mathcal{X}|).
\end{aligned}$$

514 Combining the lower bound for rewards in the two phases, we get

$$\begin{aligned}
\prod_{t=1}^T \mathbb{E}[\langle X_t, \theta^* \rangle]^\frac{1}{T} &\geq \prod_{t=1}^T \left(\mathbb{E}[\langle X_t, \theta^* \rangle \mid E_1 \cap E_2] \cdot \mathbb{P}\{E_1 \cap E_2\} \right)^\frac{1}{T} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} \right) \left(1 - 75 \sqrt{\frac{d}{T \langle x^*, \theta^* \rangle}} \log(T|\mathcal{X}|) \right) \mathbb{P}\{E_1 \cap E_2\} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} - 75 \sqrt{\frac{d}{T \langle x^*, \theta^* \rangle}} \log(T|\mathcal{X}|) \right) \mathbb{P}\{E_1 \cap E_2\} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} - 75 \sqrt{\frac{d}{T \langle x^*, \theta^* \rangle}} \log(T|\mathcal{X}|) \right) \left(1 - \frac{2 \log T}{T} \right) \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))3\sqrt{Td \log(T|\mathcal{X}|)}}{T} - 75 \sqrt{\frac{d}{T \langle x^*, \theta^* \rangle}} \log(T|\mathcal{X}|) - \frac{2 \log T}{T} \right) \\
&\geq \langle x^*, \theta^* \rangle - 75 \sqrt{\frac{\langle x^*, \theta^* \rangle d \nu}{T}} \log(T|\mathcal{X}|) - 6 \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{T}} \log(2(d+1)) \langle x^*, \theta^* \rangle.
\end{aligned}$$

515 Hence the Nash Regret can be bounded as

$$\begin{aligned}
\text{NR}_T &= \langle x^*, \theta^* \rangle - \left(\prod_{t=1}^T \mathbb{E}[\langle X_t, \theta^* \rangle] \right)^{1/T} \\
&\leq 75 \sqrt{\frac{\langle x^*, \theta^* \rangle d \nu}{T}} \log(T|\mathcal{X}|) + 6 \sqrt{\frac{d \nu \log(T|\mathcal{X}|)}{T}} \log(2(d+1)) \langle x^*, \theta^* \rangle.
\end{aligned}$$

516

□

517 C \mathcal{X} independent Nash regret

518 Instead of working with probability bounds on individual arms, we construct a confidence ellipsoid
519 around θ^* . Using the notations in Algorithm 3, we first define a new set of events for the regret
520 analysis

Algorithm 3 LINNASH (Nash Confidence Bound Algorithm for Infinite Set of Arms)

Input: Arm set \mathcal{X} and horizon of play T .

- 1: Initialize matrix $\mathbf{V} \leftarrow [0]_{d,d}$ and number of rounds $\tilde{T} = 3\sqrt{Td^{2.5}\nu \log(T)}$.
 - Part I
 - 2: Generate arm sequence \mathcal{S} for the first \tilde{T} rounds using Algorithm [1](#)
 - 3: **for** $t = 1$ to \tilde{T} **do**
 - 4: Pull the next arm X_t from the sequence \mathcal{S} .
 - 5: Observe reward r_t and update $\mathbf{V} \leftarrow \mathbf{V} + X_t X_t^T$
 - 6: **end for**
 - 7: Set estimate $\hat{\theta} := \mathbf{V}^{-1} \left(\sum_{t=1}^{\tilde{T}} r_t X_t \right)$
 - 8: Find $\eta = \max_{z \in \mathcal{X}} \langle z, \hat{\theta} \rangle$
 - 9: Update $\tilde{\mathcal{X}} \leftarrow \{x \in \mathcal{X} : \langle x, \hat{\theta} \rangle \geq \eta - 16\sqrt{\frac{3\eta d^{\frac{5}{2}} \nu \log(T)}{\tilde{T}}}\}$
 - 10: $T' \leftarrow \frac{2}{3} \tilde{T}$
 - Part II
 - 11: **while** end of time horizon T is reached **do**
 - 12: Initialize $V = [0]_{d,d}$ to be an all zeros $d \times d$ matrix and $s = [0]_d$ to be an all-zeros vector.
// Beginning of new phase.
 - 13: Find the probability distribution $\lambda \in \Delta(\tilde{\mathcal{X}})$ by maximizing the following objective

$$\log \text{Det}(\mathbf{V}(\lambda)) \text{ subject to } \lambda \in \Delta(\tilde{\mathcal{X}}) \text{ and } \text{Supp}(\lambda) \leq d(d+1)/2. \quad (27)$$
 - 14: **for** a in $\text{Supp}(\lambda)$ **do**
 - 15: Pull a for the next $\lceil \lambda_a T' \rceil$ rounds.
 - 16: Observe rewards and Update $\mathbf{V} \leftarrow \mathbf{V} + \lceil \lambda_a T' \rceil \cdot a a^T$
 - 17: Observe $\lceil \lambda_a T' \rceil$ corresponding rewards z_1, z_2, \dots and update $s \leftarrow s + (\sum_j z_j) a$.
 - 18: **end for**
 - 19: Estimate $\hat{\theta} = \mathbf{V}^{-1} \left(\sum_{t \in \mathcal{E}} r_t X_t \right)$
 - 20: Find $\eta = \max_{z \in \mathcal{X}} \langle z, \hat{\theta} \rangle$
 - 21: $\tilde{\mathcal{X}} \leftarrow \{x \in \mathcal{X} : \langle x, \hat{\theta} \rangle \geq \eta - 16\sqrt{\frac{\eta d^{\frac{5}{2}} \log(T)}{T'}}\}$
 - 22: $T' \leftarrow 2 \times T'$ // End of phase.
 - 23: **end while**
-

521 G_1 During Part I arms from the D-optimal design are chosen at least $\tilde{T}/3$ times. If $\langle x^*, \theta^* \rangle \geq$
 522 $196\sqrt{\frac{d^{2.5}}{\tilde{T}} \log T}$, then $\hat{\theta}$ calculated at the end of Part I satisfies,

$$\|\hat{\theta} - \theta^*\|_{\mathbf{V}} \leq 7\sqrt{\langle x^*, \theta^* \rangle d^{\frac{3}{2}} \nu \log T}$$

523 G_2 During Part II, for every phase, if $\langle x^*, \theta^* \rangle \geq 196\sqrt{\frac{d^{2.5}}{T}} \log T$ the estimators $\hat{\theta}$ satisfy the
 524 following

$$\|\hat{\theta} - \theta^*\|_{\mathbf{V}} \leq 7\sqrt{\langle x^*, \theta^* \rangle d^{\frac{3}{2}} \nu \log T}$$

525 C.1 Regret Analysis

526 WLOG let us assume that $\langle x^*, \theta^* \rangle \geq 196\frac{d^{1.25}}{\sqrt{T}} \log T$, otherwise the regret bound is trivially satisfied.
 527 Let \mathcal{B} denote the unit ball in \mathbb{R}^d , we have

$$\begin{aligned} \|\hat{\theta} - \theta^*\|_{\mathbf{V}} &= \left\| \mathbf{V}^{\frac{1}{2}} (\hat{\theta} - \theta^*) \right\|_2 \\ &= \max_{y \in \mathcal{B}} \langle y, \mathbf{V}^{\frac{1}{2}} (\hat{\theta} - \theta^*) \rangle \end{aligned}$$

528 We construct an ε -net for the unit ball, which we will refer to as \mathcal{C}_ε . We define $y_\varepsilon =$
 529 $\arg \min_{b \in \mathcal{B}} \|b - y\|_2$

$$\begin{aligned} \|\hat{\theta} - \theta^*\|_{\mathbf{V}} &= \max_{y \in \mathcal{B}} \langle y - y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle + \langle y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle \\ &\leq \max_{y \in \mathcal{B}} \|y - y_\varepsilon\|_2 \left\| \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \right\|_2 + \langle y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle \\ &\leq \varepsilon \left\| \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \right\|_2 + \langle y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle \end{aligned}$$

530 Rearranging we get

$$\left\| \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \right\|_{\mathbf{V}} \leq \frac{1}{1 - \varepsilon} \langle y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle \quad (28)$$

531 In the following lemmas we show that $\langle y_\varepsilon, \mathbf{V}^{\frac{1}{2}}(\hat{\theta} - \theta^*) \rangle$ is small for all values of y_ε .

Lemma 18. *Let x_1, x_2, \dots, x_n be a sequence of fixed arm pulls (from a set \mathcal{X}) such that each arm x in the support λ from D -optimal design is pulled at least $\lceil \lambda_x \tau \rceil$ times. Consider $\mathbf{V} = \sum_{j=1}^s x_j x_j^\top$ and let w be a vector such that $\|w\|_2 \leq 1$ and $\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \geq 6\sqrt{\frac{d}{\tau}} \log(\mathbb{T}|\mathcal{C}_\varepsilon|)$. Then, with probability greater than $1 - \frac{2}{\mathbb{T}|\mathcal{C}_\varepsilon|}$, we have,*

$$|\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* - \hat{\theta} \rangle| \leq \left(3\sqrt{\frac{nd}{\tau}} \log(\mathbb{T}|\mathcal{C}_\varepsilon|) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}}$$

532 *Proof.* We will make use of Lemma 5. We find the γ parameter used in the lemma. We have

$$\begin{aligned} \left(w \mathbf{V}^{\frac{1}{2}} \right)^\top \mathbf{V}^{-1} X_t &\leq \left\| w \mathbf{V}^{\frac{1}{2}} \right\|_{\mathbf{V}^{-1}} \left\| \mathbf{V}^{-1} X_t \right\|_{\mathbf{V}} \\ &\leq \|w\|_2 \|X_t\|_{\mathbf{V}^{-1}} \\ &\leq \|X_t\|_{\mathbf{V}^{-1}} \quad (\text{since } \|w\| \leq 1) \end{aligned}$$

533 Let A_λ be the optimal design matrix then we have $\mathbf{V} \succ \tau A_\lambda$. This gives us the following

$$\begin{aligned} \left(w \mathbf{V}^{\frac{1}{2}} \right)^\top \mathbf{V}^{-1} X_t &\leq \|X_t\|_{\mathbf{V}^{-1}} \\ &\leq \|X_t\|_{\frac{1}{\tau} A_\lambda^{-1}} \\ &\leq \sqrt{\frac{d}{\tau}} \quad (\text{By Theorem 2}) \end{aligned}$$

534 We use Corollary 8 with $\gamma = \sqrt{\frac{d}{\tau}}$ and $\delta = \left(6\sqrt{\frac{d}{\tau}} \frac{\nu \log(\mathbb{T}|\mathcal{C}_\varepsilon|)}{\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle} \right)^{\frac{1}{2}}$. Note that $\delta \in [0, 1]$ since

535 $\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \geq 6\sqrt{\frac{d}{\tau}} \log(\mathbb{T}|\mathcal{C}_\varepsilon|)$. We have the following probability bound

$$\begin{aligned} \mathbb{P} \left\{ |\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* - \hat{\theta} \rangle| \geq \left(6\sqrt{\frac{d}{\tau}} \nu \log(\mathbb{T}|\mathcal{C}_\varepsilon|) \langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \right)^{\frac{1}{2}} \right\} &\leq 2 \exp \left(- \frac{6\sqrt{\frac{d}{\tau}} \log(\mathbb{T}|\mathcal{C}_\varepsilon|) \langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle}{3\sqrt{\frac{d}{\tau}}} \right) \\ &\leq \frac{2}{\mathbb{T}|\mathcal{C}_\varepsilon|} \end{aligned}$$

536 We can get an upper bound on the term $\langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle$ as follows

$$\begin{aligned} \langle w \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle &\leq \|w\|_2 \left\| \mathbf{V}^{\frac{1}{2}} \theta^* \right\|_2 \\ &\leq \sqrt{\theta^{*T} \mathbf{V} \theta^*} \quad (\text{since } \|w\| \leq 1) \\ &= \sqrt{\left(\sum_{i \in [n]} \theta_i^* x_i x_i^T \theta_i^* \right)} \quad (\langle x_i, \theta^* \rangle \leq \langle x^*, \theta^* \rangle) \\ &= \sqrt{n \langle x^*, \theta^* \rangle} \end{aligned}$$

537 This proves the lemma. \square

Lemma 19. Using the same notation as Lemma 18 If $\langle w\mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \leq 6\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$

$$|\langle w\mathbf{V}^{\frac{1}{2}}, \theta^* - \hat{\theta} \rangle| \leq 12\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$$

538 *Proof.* We first use Lemma 9 to show $\langle w\mathbf{V}^{\frac{1}{2}}, \hat{\theta} \rangle \leq 12\sqrt{\frac{d}{\tau}}$ by substituting $\delta = 1$, $\alpha =$
 539 $6\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$ and $\gamma = \sqrt{\frac{d}{\tau}}$. This trivially gives us $|\langle w\mathbf{V}^{\frac{1}{2}}, \theta^* - \hat{\theta} \rangle| \leq 12\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$.
 540 Next we Lemma 10 with $\delta = 1$ and $\alpha = 6\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$ which gives $|\langle w\mathbf{V}^{\frac{1}{2}}, \theta^* - \hat{\theta} \rangle| \leq$
 541 $6\sqrt{\frac{d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$. \square

Lemma 20. If the optimal arm satisfies $\langle x^*, \theta^* \rangle \geq 196\sqrt{\frac{d^{2.5}}{\mathsf{T}}} \log \mathsf{T}$

$$\mathbb{P}\{G_1\} \geq 1 - \frac{3}{\mathsf{T}}$$

and

$$\mathbb{P}\{G_2\} \geq 1 - \frac{\log \mathsf{T}}{\mathsf{T}}$$

542 *Proof.* Recall, from (28) that we aim to get a bound on $\langle y_\varepsilon \mathbf{V}^{\frac{1}{2}}, \hat{\theta} - \theta^* \rangle$ for all possible values of
 543 y_ε . The total number of arm pulls in Part I is equal to $\tilde{\mathsf{T}}$. We will now apply Lemma 18. First, from
 544 Lemma 12 we have that the arms from the solution of the D-optimal design problem are selected
 545 (with probability greater than $1 - \frac{1}{\mathsf{T}}$) at least $\tilde{\mathsf{T}}/3$ times, that is, $\tau = \tilde{\mathsf{T}}/3$. Let us consider the case
 546 where $\langle y_\varepsilon \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \geq 6\sqrt{\frac{3d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$. Taking union bound over \mathcal{C}_ε we get that the following holds
 547 with probability greater than $1 - \frac{1}{\mathsf{T}}$

$$\left\| \hat{\theta} - \theta^* \right\|_{\mathbf{V}} \leq \frac{1}{1-\varepsilon} \langle y_\varepsilon \mathbf{V}^{\frac{1}{2}}, \hat{\theta} - \theta^* \rangle \quad (\text{From (28)})$$

$$\begin{aligned} &\leq \frac{1}{1-\varepsilon} \left(3\sqrt{\frac{\tilde{\mathsf{T}}d}{\mathsf{T}}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}} && (\text{Using Lemma 18}) \\ &\leq \frac{1}{1-\varepsilon} \left(3\sqrt{3d} \log(\mathsf{T}|\mathcal{C}_\varepsilon|) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}} \end{aligned}$$

548 Since $|\mathcal{C}_\varepsilon| \leq \left(\frac{3}{\varepsilon}\right)^d$, choosing $\varepsilon = 1/2$ gives us

$$\left\| \hat{\theta} - \theta^* \right\|_{\mathbf{V}} \leq 7 \left(d^{\frac{3}{2}} \log(\mathsf{T}) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}}$$

549 Now substituting $\tau = \mathsf{T}'/3$ in Lemma 19, if $\langle y_\varepsilon \mathbf{V}^{\frac{1}{2}}, \theta^* \rangle \leq 6\sqrt{\frac{3d}{\tau}} \log(\mathsf{T}|\mathcal{C}_\varepsilon|)$, we have

$$\begin{aligned} \left\| \hat{\theta} - \theta^* \right\|_{\mathbf{V}} &\leq \frac{1}{1-\varepsilon} \langle y_\varepsilon \mathbf{V}^{\frac{1}{2}}, \hat{\theta} - \theta^* \rangle \\ &\leq 24\sqrt{\frac{d^3}{\mathsf{T}'}} \log(\mathsf{T}) && (\text{From Lemma 19 and substituting } \varepsilon = 0.5) \\ &\leq 7 \left(d^{\frac{3}{2}} \log(\mathsf{T}) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}} \end{aligned}$$

550 The last inequality is due to the fact that $\langle x^*, \theta^* \rangle \geq 196\sqrt{\frac{d^{2.5}}{\mathsf{T}}} \log \mathsf{T}$ and $\mathsf{T}' = \tilde{\mathsf{T}}/3 \geq \sqrt{\mathsf{T}d^{2.5} \log \mathsf{T}}$.

551 Similarly, for the event G_2 , an identical use of Lemma 19 and Lemma 18 with $\tau = \mathsf{T}'$ shows that, for
 552 any fixed phase, the following holds with probability greater than $1 - \frac{1}{\mathsf{T}}$

$$\left\| \hat{\theta} - \theta^* \right\|_{\mathbf{V}} \leq 7 \left(d^{\frac{3}{2}} \log(\mathsf{T}) \langle x^*, \theta^* \rangle \right)^{\frac{1}{2}}$$

553 Taking a union bound over all phases (almost $\log \mathsf{T}$) of Part II gives us the required bound on G_2 . \square

554 **Corollary 21.** If G_1 holds, then for all $x \in \mathcal{X}$, $\hat{\theta}$ calculated at the end of Part I satisfies

$$|\langle x, \hat{\theta} \rangle - \langle x, \theta^* \rangle| \leq 7 \sqrt{\frac{3 \langle x^*, \theta^* \rangle d^{2.5} \log \mathbb{T}}{\tilde{\mathbb{T}}}}$$

555 Consider any phase ℓ in Part II. If G_2 holds, then for every arm in the surviving arm set $\tilde{\mathcal{X}}$, $\hat{\theta}$
556 calculated at the end of the phase satisfies

$$|\langle x, \hat{\theta} \rangle - \langle x, \theta^* \rangle| \leq 7 \sqrt{\frac{3 \langle x^*, \theta^* \rangle d^{2.5} \log \mathbb{T}}{2^\ell \tilde{\mathbb{T}}}}.$$

557 *Proof.* First we use Hölder's inequality

$$|\langle x, \theta^* - \hat{\theta} \rangle| \leq \|x\|_{\mathbf{V}^{-1}} \|\theta^* - \hat{\theta}\|_{\mathbf{V}}. \quad (29)$$

558 Since G_1 holds, arms from the optimal design matrix are selected at least $\tilde{\mathbb{T}}/3$ times; we have by
559 Lemma 2

$$\|x\|_{\mathbf{V}^{-1}} \leq \sqrt{\frac{3d}{\tilde{\mathbb{T}}}}.$$

560 Similarly, for every Phase in Part II with $\mathbb{T}' = 2^\ell \tilde{\mathbb{T}}/3$ we have

$$\|x\|_{\mathbf{V}^{-1}} \leq \sqrt{\frac{d}{\mathbb{T}'}}.$$

561 Finally, using bound on $\|\theta^* - \hat{\theta}\|_{\mathbf{V}}$ from events G_1 and G_2 , we get the desired result. \square

Corollary 22. If $\langle x^*, \theta^* \rangle \geq 196 \sqrt{\frac{d^{2.5}}{\mathbb{T}}} \log \mathbb{T}$

$$\frac{7}{10} \langle x^*, \theta^* \rangle \leq \max_{z \in \mathcal{X}} \langle z, \hat{\theta} \rangle \leq \frac{13}{10} \langle x^*, \theta^* \rangle$$

562 *Proof.* Since $\mathbb{T}' \geq 2\tilde{\mathbb{T}}/3$, via Lemma 21 any $\hat{\theta}$ calculated in Part I or during any phase of Part II
563 satisfies

$$|\langle x, \hat{\theta} \rangle - \langle x, \theta^* \rangle| \leq 7 \sqrt{\frac{3 \langle x^*, \theta^* \rangle d^{2.5} \log \mathbb{T}}{\tilde{\mathbb{T}}}}$$

564 We have

$$\begin{aligned} \max_{z \in \mathcal{X}} \langle z, \hat{\theta} \rangle &\geq \langle x^*, \hat{\theta} \rangle \\ &\geq \langle x^*, \theta^* \rangle - 7 \sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \log \mathbb{T}}{\tilde{\mathbb{T}}}} \\ &\geq \langle x^*, \theta^* \rangle \left(1 - 7 \sqrt{\frac{d^{2.5} \log \mathbb{T}}{\langle x^*, \theta^* \rangle \tilde{\mathbb{T}}}} \right) \\ &\geq \frac{7 \langle x^*, \theta^* \rangle}{10} \quad (\text{since } \langle x^*, \theta^* \rangle \geq 196 \sqrt{\frac{d^{2.5}}{\mathbb{T}}} \log \mathbb{T} \text{ and } \tilde{\mathbb{T}} = 3 \sqrt{\mathbb{T} d^{2.5} \nu \log(\mathbb{T})}) \end{aligned}$$

565 Similarly for any $z \in \mathcal{X}$,

$$\begin{aligned} \langle z, \hat{\theta} \rangle &\leq \langle z, \theta^* \rangle + 7 \sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \log \mathbb{T}}{\tau}} \\ &\leq \langle x^*, \theta^* \rangle \left(1 + 7 \sqrt{\frac{d^{2.5} \log \mathbb{T}}{\langle x^*, \theta^* \rangle \tau}} \right) \\ &\leq \frac{13}{10} \langle x^*, \theta^* \rangle \end{aligned}$$

566 \square

567 **Lemma 23.** *If events G_1 and G_2 hold then the optimal arm x^* always exists in the surviving set $\tilde{\mathcal{X}}$*
 568 *in every phase in Step II of Alg. 3* \square

569 *Proof.* Let $\tau = \tilde{T}/3$ for Part I and $\tau = T'$ for every phase of Part II. From Lemma 21 we have for
 570 $x \in \tilde{\mathcal{X}}$

$$\begin{aligned} \langle x^*, \hat{\theta} \rangle &\geq \langle x^*, \theta^* \rangle - 7\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \log T}{\tau}} \\ &\geq \langle x, \theta^* \rangle - 7\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \log T}{\tau}} \\ &\geq \langle x, \hat{\theta} \rangle - 14\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \log T}{\tau}} \\ &\geq \langle x, \hat{\theta} \rangle - 16\sqrt{\frac{\max_{z \in \tilde{\mathcal{X}}} \langle z, \theta^* \rangle d^{2.5} \log T}{\tau}} \end{aligned} \quad (\text{Using Lemma 22})$$

571 Hence, the best arm will never satisfy the elimination criteria in Alg. 3 \square

572 **Lemma 24.** *Given that events G_1 and G_2 hold, fix any phase index ℓ in Step II of Alg. 3. For the*
 573 *surviving set of arms $\tilde{\mathcal{X}}$ at the beginning of that phase, we will have for $\tilde{T} = \sqrt{d^{2.5} T \log(T)}$*

$$\langle x, \theta^* \rangle \geq \langle x^*, \theta^* \rangle - 26\sqrt{\frac{3d^{2.5}\nu\langle x^*, \theta^* \rangle}{2^\ell \cdot \tilde{T}}} \text{ for all } x \in \tilde{\mathcal{X}} \quad (30)$$

574 *Proof.* From the second phase onwards, if an arm is pulled in a phase with phase length parameter
 575 T' , then it was not eliminated in the previous phase with phase length parameter $\frac{T'}{2}$. Additionally,
 576 since the best arm is always present in the surviving arm set $\tilde{\mathcal{X}}$ (via Lemma 23), we have

$$\begin{aligned} \langle x, \hat{\theta} \rangle &\geq \langle x^*, \hat{\theta} \rangle - 16\sqrt{\frac{\max_{z \in \tilde{\mathcal{X}}} \langle z, \hat{\theta} \rangle d^{2.5} \nu \log(T)}{\frac{T'}{2}}} \\ &\geq \langle x^*, \hat{\theta} \rangle - 26\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5} \nu \log(T)}{T'}} \end{aligned} \quad (\text{via Lemma 22})$$

577 Substituting $T' = 2^\ell \tilde{T}/3$ in the above inequality proves the Lemma. \square

578 **Theorem 2.** *Consider the stochastic linear bandits problem over a horizon of T rounds such that at*
 579 *every round $t \in [T]$, an arm $X_t \in \mathcal{X} \subset \mathbb{R}^d$ is selected and the corresponding reward r_t is obtained*
 580 *satisfying equation 2. In this setting, Algorithm 3 achieves a Nash regret of*

$$\text{NR}_T = O\left(\frac{d^{\frac{5}{4}}(\nu\langle x^*, \theta^* \rangle)^{\frac{1}{2}}}{\sqrt{T}} \log(T)\right).$$

581 *Proof.* WLOG we assume that $\langle x^*, \theta^* \rangle \geq 192\sqrt{\frac{d\nu}{T}} \log T$, otherwise the Nash Regret bound is
 582 trivially true. For Part I, the product of expected rewards satisfies

$$\begin{aligned} \prod_{t=1}^{\tilde{T}} \mathbb{E}[\langle X_t, \theta^* \rangle \mid G_1 \cap G_2] &\geq \left(\frac{\langle x^*, \theta^* \rangle}{2(d+1)}\right)^{\frac{\tilde{T}}{2}} \quad (\text{From Lemma 4}) \\ &= \langle x^*, \theta^* \rangle^{\frac{\tilde{T}}{2}} \left(1 - \frac{1}{2}\right)^{\frac{\log(2(d+1))\tilde{T}}{2}} \\ &\geq \langle x^*, \theta^* \rangle^{\frac{\tilde{T}}{2}} \left(1 - \frac{\log(2(d+1))\tilde{T}}{T}\right) \end{aligned}$$

583 For Part II, we use Lemma [7](#). Let \mathcal{E}_i denote the time interval of i^{th} phase and let \mathbb{T}'_i be the phase
584 length parameter in that phase. Recall that $|\mathcal{E}_i| \leq \mathbb{T}'_i + \frac{d(d+1)}{2}$. Also, the algorithm runs for at most
585 $\log \mathbb{T}$ phases. Hence, we have

$$\begin{aligned}
\prod_{t=\bar{\mathbb{T}}+1}^{\mathbb{T}} \mathbb{E}[\langle X_t, \theta^* \rangle | G_1 \cap G_2]^{\frac{1}{\mathbb{T}}} &= \prod_{\mathcal{E}_j} \prod_{t \in \mathcal{E}_j} \mathbb{E}[\langle X_t, \theta^* \rangle | G_1 \cap G_2]^{\frac{1}{\mathbb{T}}} \\
&\geq \prod_{\mathcal{E}_j} \left(\langle x^*, \theta^* \rangle - 26 \sqrt{\frac{d^{2.5} \nu \langle x^*, \theta^* \rangle \log(\mathbb{T})}{\mathbb{T}'_j}} \right)^{\frac{|\mathcal{E}_j|}{\mathbb{T}}} \\
&\geq \langle x^*, \theta^* \rangle^{\frac{\mathbb{T}-\bar{\mathbb{T}}}{\mathbb{T}}} \prod_{i=1}^{\log \mathbb{T}} \left(1 - 26 \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right)^{\frac{|\mathcal{E}_j|}{\mathbb{T}}} \\
&\geq \langle x^*, \theta^* \rangle^{\frac{\mathbb{T}-\bar{\mathbb{T}}}{\mathbb{T}}} \prod_{i=1}^{\log \mathbb{T}} \left(1 - 52 \frac{|\mathcal{E}_j|}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right)
\end{aligned}$$

586 The last inequality is due to the fact that $(1-x)^r \geq (1-2rx)$ where $r \in [0, 1]$ and $x \in [0, 1/2]$.
587 Note that the term $\sqrt{\frac{d^{2.5} \log(\mathbb{T})}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \leq 1/2$ for $\langle x^*, \theta^* \rangle \geq 192 \sqrt{\frac{d^{2.5}}{\mathbb{T}} \log \mathbb{T}}$, $\mathbb{T}'_j \geq 2\sqrt{\mathbb{T} d^{2.5} \log \mathbb{T}}$ and
588 $\mathbb{T} \geq e^6$. We now further simplify the expression as shown below

$$\begin{aligned}
\prod_{j=1}^{\log \mathbb{T}} \left(1 - 52 \frac{|\mathcal{E}_j|}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right) &\geq \prod_{j=1}^{\log \mathbb{T}} \left(1 - 52 \frac{\mathbb{T}'_j + \frac{d(d+1)}{2}}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle \mathbb{T}'_j}} \right) \\
&\geq \prod_{j=1}^{\log \mathbb{T}} \left(1 - 78 \frac{\sqrt{\mathbb{T}'_j}}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle}} \right) \\
&\hspace{15em} \text{(assuming } \mathbb{T}'_j \geq d(d+1) \text{)} \\
&\geq 1 - 78 \frac{1}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle}} \left(\sum_{j=1}^{\log \mathbb{T}} \sqrt{\mathbb{T}'_j} \right) \\
&\geq 1 - 78 \frac{1}{\mathbb{T}} \sqrt{\frac{d^{2.5} \nu \log(\mathbb{T})}{\langle x^*, \theta^* \rangle}} \left(\sqrt{\mathbb{T} \log \mathbb{T}} \right) \\
&\hspace{15em} \text{(By Cauchy Schwarz)} \\
&\geq 1 - 78 \sqrt{\frac{d^{2.5} \nu}{\mathbb{T} \langle x^*, \theta^* \rangle}} \log(\mathbb{T}).
\end{aligned}$$

589 Combining the lower bound for rewards in the two phases, we get

$$\begin{aligned}
\prod_{t=1}^T \mathbb{E}[\langle X_t, \theta^* \rangle]^{\frac{1}{T}} &\geq \prod_{t=1}^T \left(\mathbb{E}[\langle X_t, \theta^* \rangle \mid G_1 \cap G_2] \cdot \mathbb{P}\{G_1 \cap G_2\} \right)^{\frac{1}{T}} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} \right) \left(1 - 78\sqrt{\frac{d^{2.5}}{T\langle x^*, \theta^* \rangle} \log(T)} \right) \mathbb{P}\{G_1 \cap G_2\} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} - 78\sqrt{\frac{d^{2.5}}{T\langle x^*, \theta^* \rangle} \log(T)} \right) \mathbb{P}\{G_1 \cap G_2\} \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))\tilde{T}}{T} - 78\sqrt{\frac{d^{2.5}}{T\langle x^*, \theta^* \rangle} \log(T)} \right) \left(1 - \frac{2 \log T}{T} \right) \\
&\geq \langle x^*, \theta^* \rangle \left(1 - \frac{\log(2(d+1))3\sqrt{Td \log(T)}}{T} - 78\sqrt{\frac{d^{2.5}}{T\langle x^*, \theta^* \rangle} \log(T)} - \frac{2 \log T}{T} \right) \\
&\geq \langle x^*, \theta^* \rangle - 78\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5}}{T} \log(T)} - 2\frac{\langle x^*, \theta^* \rangle \log(2(d+1))3\sqrt{d \log(T)}}{\sqrt{T}}.
\end{aligned}$$

590 Hence the Nash Regret can be bounded as

$$\begin{aligned}
\text{NR}_T &= \langle x^*, \theta^* \rangle - \left(\prod_{t=1}^T \mathbb{E}[\langle X_t, \theta^* \rangle] \right)^{1/T} \\
&\leq 78\sqrt{\frac{\langle x^*, \theta^* \rangle d^{2.5}}{T} \log(T)} + 2\frac{\langle x^*, \theta^* \rangle \log(2(d+1))3\sqrt{d \log(T)}}{\sqrt{T}}.
\end{aligned}$$

591

□