

Appendix

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A The Canonical Model of Bandits

We extend the general canonical model of bandits (Chapter 4, Lattimore and Szepesvári (2018)) with ϵ -global differential privacy. The canonical model with ϵ -global DP consists of a *privacy-preserving policy* π^ϵ and an *environment* ν . The policy interacts with the environment up to a given time horizon T to produce a history $\mathcal{H}_T \triangleq \{(A_t, R_t)\}_{t=1}^T$. The iterative steps of this interaction process are:

1. the probability of choosing an action $A_t = a$ at time t is dictated only by the policy $\pi_t^\epsilon(a|\mathcal{H}_{t-1})$,
2. the distribution of reward R_t is P_{A_t} and is conditionally independent of the previous observed history \mathcal{H}_{t-1} .

Let us formalise this interaction by defining an ϵ -global DP policy, the environment and the probability space produced by this interaction.

Let $T \in \mathbb{N}$ be the horizon. Let $\nu = (P_a : a \in [K])$ a bandit instance with K arms. For each $t \in [T]$, let $\Omega_t = ([K] \times \mathbb{R})^t \subset \mathbb{R}^{2t}$ and $\mathcal{F}_t = \mathfrak{B}(\Omega_t)$ with \mathfrak{B} being the Borel set.

Definition 3. A policy π is a sequence $(\pi_t)_{t=1}^T$, where π_t is a probability kernel from $(\Omega_t, \mathcal{F}_t)$ to $([K], 2^{[K]})$. Since $[K]$ is discrete, we adopt the convention that for $i \in [K]$,

$$\pi_t(i | a_1, r_1, \dots, a_{t-1}, r_{t-1}) = \pi_t(\{i\} | a_1, r_1, \dots, a_{t-1}, r_{t-1})$$

and for a sequence of actions $a^T \triangleq [a_1, \dots, a_T]$ and a sequence of rewards $r^T \triangleq [r_1, \dots, r_T]$:

$$\pi(a^T | r^T) = \prod_{t=1}^T \pi_t(a_t | a_1, r_1, \dots, a_{t-1}, r_{t-1})$$

A policy π^ϵ is ϵ -global DP, if

$$\pi^\epsilon(a^T | r^T) \leq e^\epsilon \pi^\epsilon(a^T | r'^T)$$

for every sequence of actions a^T and every two neighbouring reward streams r^T, r'^T : $\exists j \in [1, T]$ such that $r_j \neq r'_j$ and $\forall t \neq j$ $r_t = r'_t$.

Let λ be a σ -finite measure on $(\mathbb{R}, \mathfrak{B}(\mathbb{R}))$ for which P_a is absolutely continuous with respect to λ for all $a \in [K]$. Let $p_a = dP_a/d\lambda$ be the Radon–Nikodym derivative of P_a with respect to λ , which is a function $p_a : \mathbb{R} \rightarrow \mathbb{R}$ such that $\int_B p_a d\lambda = P_a(B)$ for all $B \in \mathfrak{B}(\mathbb{R})$. Letting ρ be the counting measure with $\rho(B) = |B|$, the density $p_{\nu\pi^\epsilon} : \Omega_T \rightarrow \mathbb{R}$ can now be defined with respect to the product measure $(\rho \times \lambda)^T$ by

$$p_{\nu\pi^\epsilon}(a_1, r_1, \dots, a_T, r_T) \triangleq \prod_{t=1}^T \pi_t(a_t | a_1, r_1, \dots, a_{t-1}, r_{t-1}) p_{a_t}(r_t)$$

and $\mathcal{P}_{\nu\pi^\epsilon}$ be defined by

$$\mathcal{P}_{\nu\pi^\epsilon}(B) \triangleq \int_B p_{\nu\pi^\epsilon}(\omega) (\rho \times \lambda)^T(d\omega) \quad \text{for all } B \in \mathcal{F}_T$$

Hence $(\Omega_T, \mathcal{F}_T, \mathcal{P}_{\nu\pi^\epsilon})$ is a probability space over histories induced by the interaction between π^ϵ and ν .

We define also a marginal distribution over a sequence of actions by

$$m_{\nu\pi^\epsilon}(a_1, \dots, a_T) \triangleq \int_{r_1, \dots, r_T} p_{\nu\pi^\epsilon}(a_1, r_1, \dots, a_T, r_T) dr_1 \dots dr_T,$$

and for all $C \in \mathcal{P}([K]^T)$,

$$M_{\nu\pi^\epsilon}(C) \triangleq \sum_{(a_1, \dots, a_T) \in C} m_{\nu\pi^\epsilon}(a_1, a_2, \dots, a_T).$$

Hence, $([K]^T, \mathcal{P}([K]^T), M_{\nu\pi^\epsilon})$ is a probability space over sequence of actions produced when π^ϵ interacts with ν for T time-steps.

B Distinguishing Environments with Partial Information and Global DP

In this section, we first revisit the Karwa-Vadhan Lemma (Lemma 6.1, (Karwa and Vadhan, 2017)) that bounds the multiplicative distance between marginal distributions induced by a differentially private mechanism, when the datasets are generated using two different distributions \mathbb{P} and \mathbb{Q} . We generalise this result to the setting where the inputs are not identically distributed. We call this Sequential Karwa-Vadhan Lemma (Lemma 2) and apply it to upper bound the Kullback-Leibler (KL) divergence between the marginal distributions $M_{\nu\pi^\epsilon}$ and $M_{\nu'\pi^\epsilon}$, when π^ϵ is an ϵ -global DP policy, and ν and ν' are two different environments (Theorem 10).

Karwa-Vadhan Lemma. Let \mathbb{P} and \mathbb{Q} be two distributions, and $\text{TV}(\mathbb{P} \parallel \mathbb{Q})$ be the total variation distance between these two distributions. Let \mathcal{M} be an (ϵ, δ) -differentially private mechanism that runs on the set of samples $\{x_1, \dots, x_T\}$. For any event E in \mathcal{M} 's output space, $\mathcal{M}(E|X_1 = x_1, \dots, X_T = x_T)$ denotes the probability that \mathcal{M} outputs an element in E given the input x_1, \dots, x_T , and

$$\mathbb{M}_{\mathbb{P}}(E) \triangleq \int \mathcal{M}(E|X_1, \dots, X_T) d\mathbb{P}(X_1, \dots, X_T)$$

is the marginal distribution induced by the DP mechanism when the data is generated from the distribution \mathbb{P} .

Theorem 9 (Lemma 6.1, Karwa and Vadhan (2017)). *If a mechanism \mathcal{M} satisfies (ϵ, δ) -DP, then for every event E in the output space of \mathcal{M} , the marginal distributions induced by distributions \mathbb{P} and \mathbb{Q} satisfy*

$$\mathbb{M}_{\mathbb{P}}(E) \leq e^{\epsilon'} \mathbb{M}_{\mathbb{Q}}(E) + \delta',$$

where $\epsilon' \triangleq (6\epsilon T)\text{TV}(\mathbb{P} \parallel \mathbb{Q})$ and $\delta' \triangleq (4e^{\epsilon'} T\delta)\text{TV}(\mathbb{P} \parallel \mathbb{Q})$.

We extend this result for the setting where the data is not identically distributed.

B.1 Sequential Karwa-Vadhan Lemma

Let $\{\mathbb{P}_1, \dots, \mathbb{P}_T\}$ and $\{\mathbb{Q}_1, \dots, \mathbb{Q}_T\}$ two sets of independent distributions.

Given the samples X_1, \dots, X_T generated from the distributions $\mathbb{P}_1, \dots, \mathbb{P}_T$, we define the corresponding marginal distribution induced by \mathcal{M} as

$$\mathbb{M}_{\mathbb{P}_1, \dots, \mathbb{P}_T}(E) \triangleq \int \mathcal{M}(E|X_1, \dots, X_T) d\mathbb{P}_1(X_1) d\mathbb{P}_2(X_2) \dots d\mathbb{P}_T(X_T)$$

Lemma 2 (Sequential Karwa-Vadhan Lemma). *If \mathcal{M} is a mechanism satisfying (ϵ, δ) -DP, then for every event E in the output space of \mathcal{M} , the marginal distributions induced by the two sets of independent distributions $\{\mathbb{P}_1, \dots, \mathbb{P}_T\}$ and $\{\mathbb{Q}_1, \dots, \mathbb{Q}_T\}$ satisfy*

$$\mathbb{M}_{\mathbb{P}_1, \dots, \mathbb{P}_T}(E) \leq e^{\epsilon'} \mathbb{M}_{\mathbb{Q}_1, \dots, \mathbb{Q}_T}(E) + \delta',$$

where $\epsilon' = 6\epsilon \sum_{i=1}^T \text{TV}(\mathbb{P}_i \parallel \mathbb{Q}_i)$ and $\delta' = 4e^{\epsilon'} \delta \sum_{i=1}^T \text{TV}(\mathbb{P}_i \parallel \mathbb{Q}_i)$

Proof. We extend the proof proposed by (Karwa and Vadhan, 2017) to the non-identical distribution setting. The main observation is that the proof follows naturally if the data is generated from different distributions by just adapting the coupling to the case of different distributions. For completeness, we present the whole proof with all the adapted changes.

We construct a coupling between $\bigotimes_{i=1}^T \mathbb{P}_i$ and $\bigotimes_{i=1}^T \mathbb{Q}_i$ that allows us to control the hamming distance between samples generated from this distributions.

Let us denote $p_i \triangleq \text{TV}(\mathbb{P}_i \parallel \mathbb{Q}_i)$, $F_i \triangleq \max(\mathbb{P}_i - \mathbb{Q}_i, 0)$, $G_i \triangleq \max(\mathbb{Q}_i - \mathbb{P}_i, 0)$, and $C_i \triangleq \min(\mathbb{P}_i, \mathbb{Q}_i)$. It is easy to see that $\mathbb{P}_i = F_i + C_i$ and $\mathbb{Q}_i = G_i + C_i$.

Given the aforementioned notations, we consider the following algorithm to generate $2T$ samples:

For $i = 1$ to T , generate H_i from Bernoulli(p_i)

- (a) If $H_i = 1$, sample $X_i \propto F_i$ and $X'_i \propto G_i$
- (b) If $H_i = 0$, sample $X_i \propto C_i$ and set $X'_i = X_i$.

Here $X_i \propto F_i$ means that X_i is generated from a distribution defined by normalizing F_i .

This construction satisfies the following properties:

1. $\underline{X} \triangleq (X_1, \dots, X_T) \sim \bigotimes_{i=1}^T \mathbb{P}_i \triangleq \mathbb{D}_0$.
2. $\underline{X}' \triangleq (X'_1, \dots, X'_T) \sim \bigotimes_{i=1}^T \mathbb{Q}_i \triangleq \mathbb{D}_1$.
3. $\|\underline{X} - \underline{X}'\|_{\text{Hamming}} = \sum_{i=1}^T H_i \triangleq H$.

Now, we introduce the following shorthand for the marginal distributions at step h

$$m_j(h) \triangleq \int_{\underline{x}} \mathcal{M}(E | \underline{X} = \underline{x}) d\mathbb{D}_j(\underline{X} | H = h)$$

for $j \in \{0, 1\}$ and $p(h) = \mathbb{P}(H = h)$. For $j \in \{0, 1\}$ and any event E , we have, by definition,

$$\mathbb{M}_j(E) = \sum_{h=0}^T m_j(h) p(h)$$

Fact 1: For $j \in \{0, 1\}$, $m_j(h) \leq e^\epsilon m_j(h-1) + \delta$ for $h = 1, \dots, T$, and $m_1(0) = m_0(0)$.

We defer the proof of Fact 1 to the end of this proof.

By Fact 1, for $j \in \{0, 1\}$, we have

$$m_j(h) \leq e^{h\epsilon} m_j(0) + \frac{e^{h\epsilon} - 1}{e^\epsilon - 1} \delta$$

Now, we obtain

$$\begin{aligned} \mathbb{M}_j(E) &= \sum_{h=0}^T p(h) m_j(h) \\ &= \mathbb{E}[m_j(H)] \\ &\leq \mathbb{E}[e^{H\epsilon} m_j(0) + \frac{e^{H\epsilon} - 1}{e^\epsilon - 1} \delta] \\ &= m_j(0) \cdot \mathbb{E}[e^{H\epsilon}] + \frac{\delta}{e^\epsilon - 1} \cdot (\mathbb{E}[e^{H\epsilon}] - 1) \\ &= m_j(0) \cdot \prod_{i=1}^T (1 - p_i + p_i \cdot e^\epsilon) + \frac{\delta}{e^\epsilon - 1} \cdot \left(\prod_{i=1}^T (1 - p_i + p_i \cdot e^\epsilon) - 1 \right) \end{aligned} \quad (9)$$

The last equality holds due to that fact that for any $t > 0$, $\mathbb{E}[e^{tH}] = \prod_{i=1}^T (1 - p_i + p_i \cdot e^t)$.

Similarly, we obtain

$$\mathbb{M}_j(E) \geq m_j(0) \prod_{i=1}^T (1 - p_i + p_i \cdot e^{-\epsilon}) + \frac{\delta}{e^{-\epsilon} - 1} \cdot \left(\prod_{i=1}^T (1 - p_i + p_i \cdot e^{-\epsilon}) - 1 \right) \quad (10)$$

Combining inequalities 9 and 10, we get

$$\mathbb{M}_0(E) \leq \left[\prod_{i=1}^T \left(\frac{1 - p_i + p_i \cdot e^\epsilon}{1 - p_i + p_i \cdot e^{-\epsilon}} \right) \right] \cdot \left(\mathbb{M}_1(E) + \frac{1 - \prod_{i=1}^T (1 - p_i + p_i \cdot e^{-\epsilon})}{1 - e^{-\epsilon}} \cdot \delta \right)$$

$$+ \frac{\prod_{i=1}^T (1 - p_i + p_i \cdot e^{-\epsilon}) - 1}{e^\epsilon - 1} \cdot \delta \quad (11)$$

From Lemma 6.1 of (Karwa and Vadhan, 2017), we know that

$$\log \left(\frac{1 - p_i + p_i \cdot e^\epsilon}{1 - p_i + p_i \cdot e^{-\epsilon}} \right) \leq 6\epsilon p_i,$$

Thus,

$$\prod_{i=1}^T \left(\frac{1 - p_i + p_i \cdot e^\epsilon}{1 - p_i + p_i \cdot e^{-\epsilon}} \right) \leq e^{6\epsilon \sum_{i=1}^T p_i} \triangleq e^{\epsilon'}, \quad (12)$$

and

$$e^{\epsilon'} \cdot \frac{1 - \prod_{i=1}^T (1 - p_i + p_i \cdot e^{-\epsilon})}{1 - e^{-\epsilon}} \cdot \delta + \frac{\prod_{i=1}^T (1 - p_i + p_i \cdot e^\epsilon) - 1}{e^\epsilon - 1} \cdot \delta \quad (13)$$

$$\leq e^{\epsilon'} \cdot \frac{1 - \exp(2(\sum_{i=1}^T p_i) \cdot (e^{-\epsilon} - 1))}{1 - e^{-\epsilon}} \cdot \delta + \frac{\exp(2(\sum_{i=1}^T p_i) \cdot (e^\epsilon - 1)) - 1}{e^\epsilon - 1} \cdot \delta \quad (14)$$

$$\leq e^{\epsilon'} \cdot 2 \left(\sum_{i=1}^T p_i \right) \cdot \delta + 2 \left(\sum_{i=1}^T p_i \right) \cdot \delta \quad (15)$$

$$\leq e^{\epsilon'} \cdot 4 \sum_{i=1}^T p_i \cdot \delta. \quad (16)$$

Substituting Equations (12) and (16) in Equation 11, we obtain

$$\mathbb{M}_0(E) \leq e^{\epsilon'} \mathbb{M}_1(E) + \delta',$$

where $\epsilon' = 6\epsilon(\sum_{i=1}^T p_i)$ and $\delta' = 4e^{\epsilon'}\delta(\sum_{i=1}^T p_i)$. \square

Now, we prove Fact 1.

Fact 1. For $j \in \{0, 1\}$, $m_j(h) \leq e^\epsilon m_j(h-1) + \delta$ for $h = 1, \dots, T$, and $m_1(0) = m_0(0)$.

Proof. We prove the claim for $j = 0$, the other case is similar.

First, let us introduce some notations. Fix a $(h_1, \dots, h_T) \in \{0, 1\}^T$. Let $I' \triangleq \{i : h_i = 1\}$, $J \triangleq \{i : h_i = 0\}$, and r be any fixed index in I' . Let $I = I' \setminus \{r\}$ and consider the following partition of \underline{X} into three parts:

$$\underline{X} = (\underline{X}_I, X_r, \underline{X}_J),$$

where \underline{X}_I is the vector \underline{X} specified by the indices in I . By definition of the coupling, $\underline{X}_I \sim \bigotimes_{i \in I} F_i \triangleq F_I$, $X_r \sim F_r$, $\underline{X}_J \sim \bigotimes_{i \in J} C_i \triangleq C_J$. Now, let $X'_r \sim C_r$ and

$$\underline{X}' = (\underline{X}_I, X'_r, \underline{X}_J).$$

Also, let h'_1, \dots, h'_T be the binary indicators corresponding to \underline{X}' . By construction, we have the following properties:

1. $h_i = h'_i$ for all $i \neq r$
2. $h_r = 1$ and $h'_r = 0$
3. $\sum_{i=1}^T h_i = h$ and $\sum_{i=1}^T h'_i = h - 1$
4. $\mathbb{D}_j(\underline{X} | H_1 = h_1, \dots, H_T = h_T) = \mathbb{P}_{F_I}(\underline{X}_I) \mathbb{P}_{F_r}(X_r) \mathbb{P}_{C_J}(\underline{X}_J)$
5. $\mathbb{D}_j(\underline{X}' | H_1 = h'_1, \dots, H_T = h'_T) = \mathbb{P}_{F_I}(\underline{X}_I) \mathbb{P}_{C_r}(X'_r) \mathbb{P}_{C_J}(\underline{X}_J)$

Thus, we obtain

$$\begin{aligned}
& \int_{\underline{x}} \mathcal{M}(E|\underline{X} = \underline{x}) \, d\mathbb{D}_j(\underline{X}|H_1 = h_1, \dots, H_T = h_T) \\
&= \int_{\underline{x}_I} \int_{x_r} \int_{\underline{x}_J} \mathcal{M}(E|\underline{x}_I, x_r, \underline{x}_J) \, d\mathbb{P}_{F_I}(\underline{X}_I) \, d\mathbb{P}_{F_r}(X_r) \, d\mathbb{P}_{C_J}(\underline{X}_J) \\
&\leq \int_{x'_r} \int_{\underline{x}_I} \int_{x_r} \int_{\underline{x}_J} (e^\epsilon \mathcal{M}(E|\underline{x}_I, x'_r, \underline{x}_J) + \delta) \, d\mathbb{P}_{F_I}(\underline{X}_I) \, d\mathbb{P}_{F_r}(X_r) \, d\mathbb{P}_{C_r}(X'_r) \, d\mathbb{P}_{C_J}(\underline{X}_J) \\
&\leq \int_{\underline{x}_I} \int_{x'_r} \int_{\underline{x}_J} (e^\epsilon \mathcal{M}(E|\underline{x}_I, x'_r, \underline{x}_J) + \delta) \, d\mathbb{P}_{F_I}(\underline{X}_I) \, d\mathbb{P}_{C_r}(X'_r) \, d\mathbb{P}_{C_J}(\underline{X}_J) \\
&\leq e^\epsilon \int_{\underline{x}'} \mathcal{M}(E|\underline{X} = \underline{x}') \, d\mathbb{D}_j(\underline{X}'|H_1 = h'_1, \dots, H_T = h'_T) + \delta.
\end{aligned}$$

Taking expectations on both sides with respect to (H_1, \dots, H_T) proves the claim. \square

B.2 KL-divergence Decomposition with ϵ -global DP

The Sequential Karwa-Vadhan Lemma (Lemma 2) allows us to show the maximum KL-divergence induced in the distributions of actions by a global DP policy π^ϵ . The upper bound allows us to show how different the final distributions over actions induced by a global DP policy are for two different environments. Thus, in turn, it provides an information-theoretic limit on distinguishability of two environments if π^ϵ is played.

Theorem 10 (Upper Bound on KL-divergence for Bandits with ϵ -global DP). *When an ϵ -global DP policy π^ϵ interacts with two bandit instances $\nu = (P_a : a \in [K])$ and $\nu' = (P'_a : a \in [K])$ we have:*

$$D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon}) \leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon} \left[\sum_{t=1}^T \text{TV}(P_{a_t} \parallel P'_{a_t}) \right]$$

Proof. We define the marginal over the sequence of actions induced by π^ϵ for a given environment ν as

$$m_{\nu\pi^\epsilon}(a_1, \dots, a_T) \triangleq \int_{r_1, \dots, r_T} \pi^\epsilon(a_1, \dots, a_T \mid r_1, \dots, r_T) P_{a_1} \, dr_1 \dots P_{a_T} \, dr_T$$

Since π^ϵ is ϵ -global DP, using Lemma 2, we obtain

$$\log \left(\frac{m_{\nu\pi^\epsilon}(a_1, a_2, \dots, a_T)}{m_{\nu'\pi^\epsilon}(a_1, a_2, \dots, a_T)} \right) \leq 6\epsilon \sum_{t=1}^T \text{TV}(P_{a_t} \parallel P'_{a_t})$$

for every action sequence $(a_1, \dots, a_T) \in [K]^T$.

Thus,

$$\begin{aligned}
D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon}) &= \mathbb{E}_{\nu\pi^\epsilon} \left[\log \left(\frac{m_{\nu\pi^\epsilon}(A_1, A_2, \dots, A_T)}{m_{\nu'\pi^\epsilon}(A_1, A_2, \dots, A_T)} \right) \right] \\
&\leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon} \left[\sum_{t=1}^T \text{TV}(P_{a_t} \parallel P'_{a_t}) \right]
\end{aligned}$$

\square

This lemma explicates how the distinguishability of two environments ν and ν' under π^ϵ is dictated by a joint effect of global DP, in terms of the privacy budget ϵ , and the partial information available in bandits, in terms of the total variation distance between the rewards of the arms $\mathbb{E}_{\nu\pi^\epsilon} \left[\sum_{t=1}^T \text{TV}(P_{a_t} \parallel P'_{a_t}) \right]$. We leverage this lemma further to construct the minimax and problem-dependent regret lower bounds for stochastic and linear bandits with ϵ -global DP.

C Lower Bounds on Regret: Stochastic and Linear Bandits with ϵ -global DP

In order to prove the lower bounds, we adopt the general canonical bandit model introduced in Section A. The high level idea of proving bandit lower bounds is selecting two problem instances that are similar (the policy cannot statistically distinguish between them) but conflicting (actions that may be good in one instance are not good for the other).

Under ϵ -global differential privacy, a new source of "confusion" is added to the problem, i.e. any sequence of actions induced by neighbouring reward streams must be ϵ -indistinguishable. In the canonical bandit framework, this is expressed by our Theorem 10.

In the following, we plug this upper bound on KL-divergences in the classic proofs of regret lower bounds in bandits Lattimore and Szepesvári (2018) to derive our minimax and problem-dependent regret lower bounds.

Notations. Let Π be the set of all policies, and Π^ϵ be the set of all ϵ -global DP policies.

C.1 Stochastic Bandits: Minimax Lower Bound

Theorem 2 (Minimax lower bound). *For any $K > 1$ and $T \geq K - 1$, and $\epsilon > 0$, the minimax regret of stochastic bandits with ϵ -global DP satisfies*

$$\text{Reg}_{T,\epsilon}^{\text{minimax}} \geq \max \left\{ \underbrace{\frac{1}{27} \sqrt{T(K-1)}}_{\text{without global DP}}, \underbrace{\frac{1}{131} \frac{K-1}{\epsilon}}_{\text{with } \epsilon\text{-global DP}} \right\}.$$

Proof. We denote the environment corresponding to the set of K -Gaussian reward distributions with unit variance and means $\mu \in \mathbb{R}^K$ as $\mathcal{E}_{\mathcal{N}}^K(1) \triangleq \left\{ (\mathcal{N}(\mu_i, 1))_{i=1}^K : \mu = (\mu_1, \dots, \mu_K) \in \mathbb{R}^K \right\}$.

Since $\Pi^\epsilon \subset \Pi$, we can have that

$$\text{Reg}_{T,\epsilon}^{\text{minimax}} \geq \inf_{\pi \in \Pi} \sup_{\nu \in \mathcal{E}_{\mathcal{N}}^K(1)} \text{Reg}_T(\pi, \nu) \geq \frac{1}{27} \sqrt{T(K-1)}$$

The second inequality is due to Theorem 15.2 in (Lattimore and Szepesvári, 2018).

Step 1: Choosing the ‘Hard-to-distinguish’ Environments. First, we fix a policy π^ϵ in Π^ϵ .

Let Δ be a constant (to be specified later), and ν be a Gaussian bandit instance with unit variance and mean vector $\mu = (\Delta, 0, 0, \dots, 0)$.

To choose the second bandit instance, let $i \triangleq \arg \min_{a>1} \mathbb{E}_{\nu, \pi^\epsilon} [N_a(T)]$ be the least played arm in expectation other than the optimal arm 1.

The second environment ν' is then chosen to be a Gaussian bandit instance with unit variance and mean vector $\mu' = (\Delta, 0, 0, \dots, 0, 2\Delta, 0, \dots, 0)$, where $\mu'_j = \mu_j$ for every j except for $\mu'_i = 2\Delta$.

The first arm is optimal in ν and the arm i is optimal in ν' .

Since $T = \mathbb{E}_{\nu, \pi^\epsilon} [N_1(T)] + \sum_{a>1} \mathbb{E}_{\nu, \pi^\epsilon} [N_a(T)] \geq (K-1) \mathbb{E}_{\nu, \pi^\epsilon} [N_i(T)]$, we observe that

$$\mathbb{E}_{\nu, \pi^\epsilon} [N_i(T)] \leq \frac{T}{K-1}$$

Step 2: From Lower Bounding Regret to Upper Bounding KL-divergence. Now by the classic regret decomposition and Markov Inequality 6, we get⁷

$$\text{Reg}_T(\pi^\epsilon, \nu) = (T - \mathbb{E}_{\nu, \pi^\epsilon} [N_1(T)]) \Delta \geq \mathbb{M}_{\nu, \pi^\epsilon} (N_1(T) \leq T/2) \frac{T\Delta}{2},$$

and

$$\text{Reg}_T(\pi^\epsilon, \nu') = \Delta \mathbb{E}_{\nu', \pi^\epsilon} [N_1(T)] + \sum_{a \notin \{1, i\}} 2\Delta \mathbb{E}_{\nu', \pi^\epsilon} [N_a(T)] \geq \mathbb{M}_{\nu', \pi^\epsilon} (N_1(T) > T/2) \frac{T\Delta}{2}.$$

⁷In all regret lower bound proofs, we are under the probability space over sequence of actions, produced when π^ϵ interacts with ν for T time-steps. We do this to use the KL-divergence decomposition of $\mathbb{M}_{\nu, \pi^\epsilon}$

Let us define the event $A \triangleq \{N_1(T) \leq T/2\} = \{(a_1, a_2, \dots, a_T) : \text{card}(\{j : a_j = 1\}) \leq T/2\}$.

By applying the Bretagnolle–Huber inequality, we have:

$$\begin{aligned} \text{Reg}_T(\pi^\epsilon, \nu) + \text{Reg}_T(\pi^\epsilon, \nu') &\geq \frac{T\Delta}{2} (M_{\nu\pi^\epsilon}(A) + M_{\nu'\pi^\epsilon}(A^c)) \\ &\geq \frac{T\Delta}{4} \exp(-D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon})) \end{aligned}$$

Step 3: KL-divergence Decomposition with ϵ -global DP. Now, we apply Theorem 10 to upper-bound the KL-Divergence between the marginals.

$$\begin{aligned} D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon}) &\leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon} \left[\sum_{t=1}^T \text{TV}(P_{a_t} \parallel P'_{a_t}) \right] \\ &\leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon} [N_i(T)] \text{TV}(p_i \parallel p'_i) \end{aligned}$$

since ν and ν' only differ in the arm i .

Finally, using Pinsker's Inequality 9, we obtain

$$\text{TV}(p_i \parallel p'_i) \leq \sqrt{\frac{1}{2} D_{\text{KL}}(\mathcal{N}(0, 1) \parallel \mathcal{N}(2\Delta, 1))} = \Delta$$

Step 4: Choosing the Worst Δ . Plugging back in the regret expression, we find

$$\begin{aligned} \text{Reg}_T(\pi^\epsilon, \nu) + \text{Reg}_T(\pi^\epsilon, \nu') &\geq \frac{T\Delta}{4} \exp(-6\epsilon \mathbb{E}_{\nu\pi^\epsilon} [N_i(T)] \Delta) \\ &\geq \frac{T\Delta}{4} \exp\left(-\frac{6\epsilon T\Delta}{K-1}\right) \end{aligned}$$

By optimising for Δ , we choose $\Delta = \frac{K-1}{6\epsilon T}$.

We conclude the proof by lower bounding $\exp(-1)$ with $\frac{48}{131}$, and using $2 \max(a, b) \geq a + b$. \square

C.2 Stochastic Bandits: Problem-dependent Lower Bound

Theorem 3 (Problem-dependent Regret Lower Bound). *Let the environment \mathcal{E} be a set of K reward distributions with finite means and a policy $\pi^\epsilon \in \Pi_{\text{cons}}(\mathcal{E}) \cap \Pi^\epsilon$ be a consistent policy⁸ over \mathcal{E} satisfying ϵ -global DP. Then, for all $\nu = (P_i)_{i=1}^K \in \mathcal{E}$, it holds that*

$$\liminf_{T \rightarrow \infty} \frac{\text{Reg}_T(\pi^\epsilon, \nu)}{\log(T)} \geq \sum_{a: \Delta_a > 0} \frac{\Delta_a}{\min\left(\underbrace{d_{\text{inf}}(P_a, \mu^*, \mathcal{M}_a)}_{\text{without global DP}}, \underbrace{6\epsilon t_{\text{inf}}(P_a, \mu^*, \mathcal{M}_a)}_{\text{with } \epsilon\text{-global DP}}\right)}.$$

Proof. Let μ_a be the mean of the a -th arm in ν , $t_a = t_{\text{inf}}(P_a, \mu^*, \mathcal{M}_a)$ and $\pi^\epsilon \in \Pi_{\text{cons}}(\mathcal{E}) \cap \Pi^\epsilon$.

Since π^ϵ is consistent, by (Theorem 16.2, Lattimore and Szepesvári (2018)), it holds that

$$\liminf_{T \rightarrow \infty} \frac{\text{Reg}_T(\pi^\epsilon, \nu)}{\log(T)} \geq \sum_{a: \Delta_a > 0} \frac{\Delta_a}{d_{\text{inf}}(P_a, \mu^*, \mathcal{M}_a)}.$$

The theorem will follow by showing, for every suboptimal arm a :

$$\liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\nu\pi^\epsilon} [N_a(T)]}{\log(T)} \geq \frac{1}{6\epsilon t_a}$$

Fix a suboptimal arm a , and let $\alpha > 0$ be an arbitrary constant.

⁸A policy π is called consistent over a class of bandits \mathcal{E} if for all $\nu \in \mathcal{E}$ and $p > 0$, it holds that $\lim_{T \rightarrow \infty} \frac{R_T(\pi, \nu)}{T^p} = 0$. We denote the class of consistent policies over a set of environments \mathcal{E} as $\Pi_{\text{cons}}(\mathcal{E})$.

Step 1: Choosing the ‘Hard-to-distinguish’ Environment. Let $\nu' \triangleq (P'_j)_{j=1}^K \in \mathcal{E}$ be a bandit with $P'_j = P_j$ for $j \neq a$ and $P'_a \in \mathcal{M}_a$ be such that $\text{TV}(P_a \parallel P'_a) \leq t_a + \alpha$ and $\mu(P'_a) > \mu^*$, which exists by the definition of t_a . Let $\mu' \in \mathbb{R}^K$ be the vector of means of distributions of ν' .

Step 2: From Lower Bounding Regret to Upper Bounding KL-divergence. For simplicity of notations, we use $\text{Reg}_T = \text{Reg}_T(\pi^\epsilon, \nu)$, $\text{Reg}'_T = \text{Reg}_T(\pi^\epsilon, \nu')$, and $A = \{(a_1, a_2, \dots, a_T) : \text{card}(\{j : a_j = 1\}) \leq T/2\}$.

Then, by regret decomposition and Markov Inequality 6, we obtain

$$\begin{aligned} \text{Reg}_T + \text{Reg}'_T &\geq \frac{T}{2} (M_{\nu\pi^\epsilon}(A)\Delta_a + M_{\nu'\pi^\epsilon}(A^c)(\mu'_a - \mu^*)) \\ &\geq \frac{T}{2} \min\{\Delta_a, \mu'_a - \mu^*\} (M_{\nu\pi^\epsilon}(A) + M_{\nu'\pi^\epsilon}(A^c)) \\ &\geq \frac{T}{4} \min\{\Delta_a, \mu'_a - \mu^*\} \exp(-D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon})) \end{aligned} \quad (17)$$

Step 3: KL-divergence Decomposition with ϵ -global DP. By Theorem 10 and the construction of the ‘hard-to-distinguish’ environments, we obtain

$$\begin{aligned} D_{\text{KL}}(M_{\nu\pi^\epsilon} \parallel M_{\nu'\pi^\epsilon}) &\leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon}[N_a(T)] \text{TV}(P_a \parallel P'_a) \\ &\leq 6\epsilon \mathbb{E}_{\nu\pi^\epsilon}[N_a(T)] (t_a + \alpha) \end{aligned}$$

Step 4: Rearranging and taking the limit inferior. Thus, we get

$$\text{Reg}_T + \text{Reg}'_T \geq \frac{T}{4} \min\{\Delta_a, \mu'_a - \mu^*\} \exp(-6\epsilon \mathbb{E}_{\nu\pi^\epsilon}[N_a(T)] (t_a + \alpha))$$

Now, taking the limit inferior on both sides leads to

$$\begin{aligned} \liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\nu\pi^\epsilon}[N_a(T)]}{\log(T)} &\geq \frac{1}{6\epsilon(t_a + \alpha)} \liminf_{T \rightarrow \infty} \frac{\log\left(\frac{T \min\{\Delta_a, \mu'_a - \mu^*\}}{4(\text{Reg}_T + \text{Reg}'_T)}\right)}{\log(T)} \\ &= \frac{1}{6\epsilon(t_a + \alpha)} \left(1 - \limsup_{T \rightarrow \infty} \frac{\log(\text{Reg}_T + \text{Reg}'_T)}{\log(T)}\right) = \frac{1}{6\epsilon(t_a + \alpha)}. \end{aligned}$$

The last equality follows from the definition of consistency, which says that for any $p > 0$, there exists a constant C_p such that for sufficiently large T , $\text{Reg}_T + \text{Reg}'_T \leq C_p T^p$. This property implies that

$$\limsup_{T \rightarrow \infty} \frac{\log(\text{Reg}_T + \text{Reg}'_T)}{\log(T)} \leq \limsup_{T \rightarrow \infty} \frac{p \log(T) + \log(C_p)}{\log(T)} = p,$$

which gives the result since $p > 0$ was an arbitrary constant.

We arrive at the claimed result by taking the limit as α tends to zero. □

Remark 2. For Bernoulli distributions, t_a is equal to Δ_a , so the private lower bound simplifies to:

$$O\left(\sum_{a: \Delta_a > 0} \Delta_a \frac{1}{\epsilon \Delta_a} \log(T)\right) = O\left(\frac{K \log(T)}{\epsilon}\right)$$

Thus, our problem-dependent regret lower bound retrieves as a special case the lower bound found in (Shariff and Sheffet, 2018) and established for Bernoulli distributions of rewards.

C.3 Stochastic Linear Bandits: Minimax Lower Bound

Theorem 4 (Minimax Regret Lower Bound). *Let $\mathcal{A} = [-1, 1]^d$ and $\Theta = \mathbb{R}^d$. Then, for any ϵ -global DP policy, we have that*

$$\text{Reg}_T^{\text{minimax}}(\mathcal{A}, \Theta) \geq \max \left\{ \underbrace{\frac{\exp(-2)}{8} d \sqrt{T}}_{\text{without global DP}}, \underbrace{\frac{\exp(-6)}{4} \frac{d}{\epsilon}}_{\text{with } \epsilon\text{-global DP}} \right\}.$$

Proof. Due to Theorem 24.1, (Lattimore and Szepesvári, 2018), it holds that,

$$\text{Reg}_T^{\text{minimax}}(\mathcal{A}, \Theta) \geq \exp(-2) \frac{d}{8} \sqrt{T}.$$

Now, we focus on proving the ϵ -global DP part of the lower bound.

Let $\Theta = \left\{ -\frac{1}{\epsilon T}, \frac{1}{\epsilon T} \right\}^d$. For $\theta, \theta' \in \Theta$, let ν and ν' be the bandit instances corresponding resp. to θ and θ' . We denote $\mathbb{M}_\theta = \mathbb{M}_{\nu, \pi^\epsilon}$ and $\mathbb{M}_{\theta'} = \mathbb{M}_{\nu', \pi^\epsilon}$. Let \mathbb{E}_θ and $\mathbb{E}_{\theta'}$ the expectations under \mathbb{M}_θ and $\mathbb{M}_{\theta'}$ respectively.

Step 1: From Lower Bounding Regret to Upper Bounding KL-divergence We begin with

$$\begin{aligned} \text{Reg}_T(\mathcal{A}, \theta) &= \mathbb{E}_\theta \left[\sum_{t=1}^T \sum_{i=1}^d (\text{sign}(\theta_i) - A_{ti}) \theta_i \right] \\ &\geq \frac{1}{\epsilon T} \sum_{i=1}^d \mathbb{E}_\theta \left[\sum_{t=1}^T \mathbb{I} \{ \text{sign}(A_{ti}) \neq \text{sign}(\theta_i) \} \right] \\ &\geq \frac{1}{\epsilon} \sum_{i=1}^d \mathbb{M}_\theta \left(\sum_{t=1}^T \mathbb{I} \{ \text{sign}(A_{ti}) \neq \text{sign}(\theta_i) \} \geq T/2 \right) \end{aligned}$$

In this derivation, the first equality holds because the optimal action satisfies $a_i^* = \text{sign}(\theta_i)$ for $i \in [d]$. The first inequality follows from an observation that $(\text{sign}(\theta_i) - A_{ti}) \theta_i \geq |\theta_i| \mathbb{I} \{ \text{sign}(A_{ti}) \neq \text{sign}(\theta_i) \}$. The last inequality is a direct application of Markov's inequality 6.

For $i \in [d]$ and $\theta \in \Theta$, we define

$$p_{\theta, i} \triangleq \mathbb{M}_\theta \left(\sum_{t=1}^T \mathbb{I} \{ \text{sign}(A_{ti}) \neq \text{sign}(\theta_i) \} \geq T/2 \right).$$

Now, let $i \in [d]$ and $\theta \in \Theta$ be fixed. Also, let $\theta'_j = \theta_j$ for $j \neq i$ and $\theta'_i = -\theta_i$. Then, by the Bretagnolle-Huber inequality,

$$p_{\theta, i} + p_{\theta', i} \geq \frac{1}{2} \exp(-D_{\text{KL}}(\mathbb{M}_\theta \parallel \mathbb{M}_{\theta'})).$$

Step 2: KL-divergence Decomposition with ϵ -global DP. From Theorem 10, we obtain that

$$\begin{aligned} D_{\text{KL}}(\mathbb{M}_\theta \parallel \mathbb{M}_{\theta'}) &\leq 6\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T \text{TV}(\mathcal{N}(\langle A_t, \theta \rangle, 1) \parallel \mathcal{N}(\langle A_t, \theta' \rangle, 1)) \right] \\ &\leq 6\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T \sqrt{\frac{1}{2} D_{\text{KL}}(\mathcal{N}(\langle A_t, \theta \rangle, 1) \parallel \mathcal{N}(\langle A_t, \theta' \rangle, 1))} \right] \\ &= 6\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T \sqrt{\frac{1}{4} [\langle A_t, \theta - \theta' \rangle]^2} \right] \\ &= 3\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T |\langle A_t, \theta - \theta' \rangle| \right] \end{aligned} \tag{18}$$

$$\begin{aligned}
&= 3\epsilon \mathbb{E}_{\nu\pi^\epsilon} \left[\sum_{t=1}^T |A_{t,i}| (2|\theta_i|) \right] \\
&\leq 3\epsilon \mathbb{E}_{\nu\pi^\epsilon} \left[T \times 2 \frac{1}{\epsilon T} \right] = 6
\end{aligned} \tag{19}$$

Here, the second inequality is a consequence of Pinsker's inequality (Lemma 9). The last inequality holds true because $A_t \in [-1, 1]^d$ and $\theta, \theta' \in \{-\frac{1}{\epsilon T}, \frac{1}{\epsilon T}\}^d$

Step 3: Choosing the 'Hard-to-distinguish' θ . We already have that

$$p_{\theta,i} + p_{\theta',i} \geq \frac{1}{2} \exp(-6)$$

Now, we apply an 'averaging hammer' over all $\theta \in \Theta$, such that $|\Theta| = 2^d$, to obtain

$$\sum_{\theta \in \Theta} \frac{1}{|\Theta|} \sum_{i=1}^d p_{\theta,i} = \frac{1}{|\Theta|} \sum_{i=1}^d \sum_{\theta \in \Theta} p_{\theta,i} \geq \frac{d}{4} \exp(-6).$$

This implies that there exists a $\theta \in \Theta$ such that $\sum_{i=1}^d p_{\theta,i} \geq d \exp(-6)/4$.

Step 4: Plugging Back θ in the Regret Decomposition. With this choice of θ , we conclude that

$$\begin{aligned}
\text{Reg}_T(\mathcal{A}, \theta) &\geq \frac{1}{\epsilon} \sum_{i=1}^d p_{\theta,i} \\
&\geq \frac{\exp(-6)}{4} \frac{d}{\epsilon}
\end{aligned}$$

□

C.4 Stochastic Linear Bandits: Problem-dependent Lower Bound

Theorem 5 (Problem-dependent Regret Lower Bound). *Let $\mathcal{A} \subset \mathbb{R}^d$ be a finite set spanning \mathbb{R}^d and $\theta \in \mathbb{R}^d$ be such that there is a unique optimal action. Then, any consistent and ϵ -global DP bandit algorithm π^ϵ satisfies*

$$\liminf_{T \rightarrow \infty} \frac{\text{Reg}_T(\mathcal{A}, \theta)}{\log(T)} \geq c(\mathcal{A}, \theta),$$

where the structural distinguishability gap is the solution of a constraint optimisation

$$c(\mathcal{A}, \theta) \triangleq \inf_{\alpha \in [0, \infty)^{\mathcal{A}}} \sum_{a \in \mathcal{A}} \alpha(a) \Delta_a, \text{ such that } \|a\|_{H_\alpha}^2 \leq \min \left\{ \underbrace{0.5 \Delta_a^2}_{\text{without global DP}}, \underbrace{3\epsilon \rho_a(\mathcal{A}) \Delta_a}_{\text{with } \epsilon\text{-global DP}} \right\}$$

for all $a \in \mathcal{A}$ with $\Delta_a > 0$, $H_\alpha = \sum_{a \in \mathcal{A}} \alpha(a) a a^\top$, and an arm-structure dependent constant $\rho_a(\mathcal{A})$.

Proof. Let $a^* = \arg\max_{a \in \mathcal{A}} \langle a, \theta \rangle$ be the optimal action, which we assumed to be unique.

By Theorem 25.1, Lattimore and Szepesvári (2018),

$$\limsup_{T \rightarrow \infty} \log(T) \|a - a^*\|_{G_T^{-1}}^2 \leq \frac{1}{2} \Delta_a^2. \tag{20}$$

Let \mathbb{M} and \mathbb{M}' be the measures on the sequence of outcomes A_1, \dots, A_T induced by θ and θ' respectively. Let $\mathbb{E}[\cdot]$ and $\mathbb{E}'[\cdot]$ be the expectation operators of \mathbb{M} and \mathbb{M}' , respectively.

Step 1: Choosing the 'Hard to distinguish' θ' . Let $\theta' \in \mathbb{R}^d$ be an alternative parameter to be chosen subsequently. We follow the usual plan of choosing θ' to be close to θ , but also ensuring that the optimal action in the bandit determined by θ' is not a^* . Let $\Delta_{\min} = \min \{\Delta_a : a \in \mathcal{A}, \Delta_a > 0\}$, $\alpha \in (0, \Delta_{\min})$ and H be a positive definite matrix (to be chosen later) such that $\|a - a^*\|_H^2 > 0$.

Given this setting, we define

$$\theta' \triangleq \theta + \frac{\Delta_a + \alpha}{\|a - a^*\|_H^2} H(a - a^*),$$

which is chosen such that $\langle a - a^*, \theta' \rangle = \langle a - a^*, \theta \rangle + \Delta_a + \alpha = \alpha$.

This means that a^* is α -suboptimal for the environment corresponding to θ' .

Step 2: From Lower Bounding Regret to Upper Bounding KL-divergence. For simplicity, we abbreviate $\text{Reg}_T = \text{Reg}_T(\mathcal{A}, \theta)$ and $\text{Reg}'_T = \text{Reg}_T(\mathcal{A}, \theta')$.

Then, by applying the classic regret decomposition and Markov's inequality 6, we obtain

$$\text{Reg}_T = \mathbb{E} \left[\sum_{a \in \mathcal{A}} N_a(T) \Delta_a \right] \geq \frac{T \Delta_{\min}}{2} \mathbb{M}(N_{a^*}(T) < T/2) \geq \frac{T \alpha}{2} \mathbb{M}(N_{a^*}(T) < T/2),$$

Since a^* is α -suboptimal in bandit θ' , it implies that

$$\text{Reg}'_T \geq \frac{T \alpha}{2} \mathbb{M}'(N_{a^*}(T) \geq T/2).$$

Now, Bretagnolle–Huber inequality implies that

$$\begin{aligned} \text{Reg}_T + \text{Reg}'_T &\geq \frac{T \alpha}{2} (\mathbb{M}(N_{a^*}(T) < T/2) + \mathbb{M}'(N_{a^*}(T) \geq T/2)) \\ &\geq \frac{T \alpha}{4} \exp(-D_{\text{KL}}(\mathbb{M} \parallel \mathbb{M}')) \end{aligned}$$

Step 3: KL-divergence Decomposition with ϵ -global DP. By Equation 18, we have that

$$\begin{aligned} D_{\text{KL}}(\mathbb{M} \parallel \mathbb{M}') &\leq 3\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T |\langle A_t, \theta - \theta' \rangle| \right] \\ &= 3\epsilon \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T \left| \left\langle A_t, \frac{\Delta_a + \alpha}{\|a - a^*\|_H^2} H(a - a^*) \right\rangle \right| \right] \\ &= 3\epsilon \frac{\Delta_a + \alpha}{\|a - a^*\|_{\bar{G}_T^{-1}}^2} \rho_T(H), \end{aligned}$$

where we define

$$\rho_T(H) \triangleq \frac{\|a - a^*\|_{\bar{G}_T^{-1}}^2}{\|a - a^*\|_H^2} \mathbb{E}_{\nu \pi^\epsilon} \left[\sum_{t=1}^T |\langle A_t, H(a - a^*) \rangle| \right]$$

Thus, after re-arrangement, we get

$$\frac{3\epsilon(\Delta_a + \alpha)}{\log(T) \|a - a^*\|_{\bar{G}_T^{-1}}^2} \rho_T(H) \geq 1 - \frac{\log((4R_T + 4R'_T)/\alpha)}{\log(T)}. \quad (21)$$

Step 4: Choosing H and Taking the Limit. The definition of consistency means that Reg_T and Reg'_T are both sub-linear in T . This implies that the second term in Equation (21) tends to zero for large T . Thus, by tending T to ∞ and α to zero, we obtain

$$\liminf_{T \rightarrow \infty} \frac{\rho_T(H)}{\log(T) \|a - a^*\|_{\bar{G}_T^{-1}}^2} \geq \frac{1}{3\epsilon \Delta_a}.$$

We now choose H to be a cluster point of the sequence $(\bar{G}_T^{-1} / \|\bar{G}_T^{-1}\|)_{T \in S}$ where $\|\bar{G}_T^{-1}\|$ is the spectral norm of the matrix \bar{G}_T^{-1} .

Fact 2: For this choice of H ,

$$\liminf_{T \rightarrow \infty} \rho_T(H) \leq \rho_a(\mathcal{A}),$$

where

$$\rho_a(\mathcal{A}) \triangleq \sum_{j=1, \|a_j\| \neq 0}^K \frac{|a_j^T(a - a^*)|}{\|a_j\|^2}.$$

Finally,

$$\limsup_{T \rightarrow \infty} \log(T) \|a - a^*\|_{\bar{G}_T^{-1}}^2 \leq 3\epsilon \Delta_a \rho_a(\mathcal{A}).$$

Combined with Equation 20, we get that

$$\limsup_{T \rightarrow \infty} \log(T) \|a - a^*\|_{\bar{G}_T^{-1}}^2 \leq \min\left(\frac{1}{2}\Delta_a^2, 3\epsilon \Delta_a \rho_a(\mathcal{A})\right).$$

Using that

$$\lim_{T \rightarrow \infty} \frac{\|a - a^*\|_{\bar{G}_T^{-1}}}{\|a\|_{\bar{G}_T^{-1}}} = 1$$

from Theorem 25.1, Lattimore and Szepesvári (2018), we get that

$$\limsup_{T \rightarrow \infty} \log(T) \|a\|_{\bar{G}_T^{-1}}^2 \leq \min\left(\frac{1}{2}\Delta_a^2, 3\epsilon \Delta_a \rho_a(\mathcal{A})\right).$$

Step 5: Getting Back to the Regret. We conclude using the same steps as in the Corollary 2 (Lattimore and Szepesvari, 2017). \square

Now, we prove Fact 2.

Fact 2. If H is a cluster point of the sequence $(\bar{G}_T^{-1} / \|\bar{G}_T^{-1}\|)_{T \in S}$ and $\|\bar{G}_T^{-1}\|$ is the spectral norm of the matrix \bar{G}_T^{-1} , then the following inequality holds true:

$$\liminf_{T \rightarrow \infty} \rho_T(H) \leq \rho_a(\mathcal{A}),$$

where

$$\rho_a(\mathcal{A}) \triangleq \sum_{j=1, \|a_j\| \neq 0}^K \frac{|a_j^T(a - a^*)|}{\|a_j\|^2}.$$

Proof. We let S be a subset so that $\bar{G}_T^{-1} / \|\bar{G}_T^{-1}\|$ converges to H on $T \in S$. Then,

$$\begin{aligned} \liminf_{T \rightarrow \infty} \rho_T(H) &\leq \liminf_{T \in S} \rho_T(\bar{G}_T^{-1} / \|\bar{G}_T^{-1}\|) \\ &= \liminf_{T \in S} \mathbb{E}_\theta \left[\sum_{t=1}^T |\langle A_t, \bar{G}_T^{-1}(a - a^*) \rangle| \right] \\ &= \liminf_{T \in S} \sum_{j=1}^K \mathbb{E}_\theta(N_j(T)) |a_j^T \bar{G}_T^{-1}(a - a^*)| \\ &= \liminf_{T \in S} \sum_{j=1, \|a_j\| \neq 0}^K \mathbb{E}_\theta(N_j(T)) |a_j^T \bar{G}_T^{-1}(a - a^*)| \end{aligned}$$

Let j be such that $\|a_j\| \neq 0$.

Now, we aim to upper bound the term $|a_j^T \bar{G}_T^{-1} (a - a^*)|$

First, we decompose $a - a^*$ into two orthogonal components, which are aligned and orthogonal to a_j respectively.

$$a - a^* = \alpha_j a_j + b_j,$$

where $a_j^\top b_j = 0$ and $\alpha_j = \frac{a_j^T (a - a^*)}{\|a_j\|^2}$.

On the other hand, we have that

$$\bar{G}_T = \mathbb{E}_\theta \left[\sum_{t=1}^T A_t A_t^\top \right] = \sum_{j=1}^K \mathbb{E}_\theta(N_j(T)) a_j a_j^\top \succeq \mathbb{E}_\theta(N_j(T)) a_j a_j^\top$$

Since

$$\left(\mathbb{E}_\theta(N_j(T)) a_j a_j^\top \right)^\dagger = \frac{1}{\mathbb{E}_\theta(N_j(T)) (a_j^\top a_j)^2} a_j a_j^\top,$$

and

$$\left(\mathbb{E}_\theta(N_j(T)) a_j a_j^\top \right)^\dagger b_j = 0,$$

only the component of $a - a^*$ in the direction of a_j matters in the dot product $a_j^T \bar{G}_T^{-1} (a - a^*)$. Thus,

$$\begin{aligned} |a_j^T \bar{G}_T^{-1} (a - a^*)| &\leq \frac{|\alpha_j|}{\mathbb{E}_\theta(N_j(T)) (a_j^\top a_j)^2} a_j^T a_j a_j^T a_j \\ &= \frac{|\alpha_j|}{\mathbb{E}_\theta(N_j(T))} \end{aligned}$$

Consequently,

$$\liminf_{T \rightarrow \infty} \rho_T(H) \leq \sum_{j=1, \|a_j\| \neq 0}^K \frac{|a_j^T (a - a^*)|}{\|a_j\|^2} \triangleq \rho_a(\mathcal{A})$$

□

Example 3 ($\rho_a(\mathcal{A})$ for an orthogonal set of arms). *If the action space is the orthogonal basis, then $\rho_a(\mathcal{A}) = 2$, because:*

$$\bar{G}_T = \begin{bmatrix} \mathbb{E}(N_1(T)) & & \\ & \ddots & \\ & & \mathbb{E}(N_d(T)) \end{bmatrix}$$

and:

$$|\langle A_t, \bar{G}_T^{-1} (a - a^*) \rangle| = \frac{1}{\mathbb{E}(N_a(T))} \mathbb{I}_{A_t=a} + \frac{1}{\mathbb{E}(N_{a^*}(T))} \mathbb{I}_{A_t=a^*}$$

so:

$$\mathbb{E} \left[\sum_{t=1}^T |\langle A_t, \bar{G}_T^{-1} (a - a^*) \rangle| \right] = 2$$

D Privacy Analysis of Algorithm 1

In this section, we prove that any bandit algorithm designed using the framework of Algorithm 1 satisfies ϵ -global DP. We establish the claim by proving ϵ -global DP for the set of private indices computed in Algorithm 1 and the final result is a consequence of the post-processing property of DP (Lemma 5).

Lemma 1 (Privacy of the $(l + 1)$ -means Computed in Algorithm 1). *Let us define the private empirical mean of the rewards between steps i and j ($i < j$) as*

$$f^\epsilon\{r_i, \dots, r_j\} \triangleq \frac{1}{j-i} \sum_{t=i}^j r_t + \text{Lap}\left(\frac{1}{(j-i)\epsilon}\right). \quad (22)$$

If $1 < t_1 < \dots < t_\ell < T$ and $r_t \in [0, 1]$, the mechanism g^ϵ mapping the sequence of rewards $(r_1, r_2, \dots, r_{T-1}, r_T)$ to $(l + 1)$ -private empirical means $(f^\epsilon\{r_1, \dots, r_{t_1-1}\}, f^\epsilon\{r_{t_1}, \dots, r_{t_2-1}\}, \dots, f^\epsilon\{r_{t_{\ell-1}}, \dots, r_{t_\ell-1}\}, f^\epsilon\{r_{t_\ell}, \dots, r_T\})$ satisfies ϵ -DP.

Proof. Let $r^T \triangleq (r_1, \dots, r_T)$ and $r'^T \triangleq (r'_1, \dots, r'_T)$ be two neighbouring reward sequences in $[0, 1]$. This implies that $\exists j \in [1, T]$ such that $r_j \neq r'_j$ and $\forall t \neq j, r_t = r'_t$.

Let ℓ' be such that $t_{\ell'} \leq j \leq t_{\ell'+1} - 1$, and follows the convention that $t_0 = 1$ and $t_{\ell+1} = T + 1$.

Let $\mu \triangleq (\mu_0, \dots, \mu_\ell)$ a fixed sequence of outcomes obtained using Equation (22). Then,

$$\frac{\mathbb{P}(g^\epsilon(r^T) = \mu)}{\mathbb{P}(g^\epsilon(r'^T) = \mu)} = \frac{\mathbb{P}\left(f^\epsilon\{r_{t_{\ell'}}, \dots, r_{t_{\ell'+1}-1}\} = \mu_{\ell'}\right)}{\mathbb{P}\left(f^\epsilon\{r_{t_{\ell'}}, \dots, r_{t_{\ell'+1}-1}\} = \mu_{\ell'}\right)} \leq e^\epsilon,$$

where the last inequality holds true because f^ϵ satisfies ϵ -DP following Theorem 1. \square

Theorem 6 (ϵ -global DP for Algorithm 1). *For any index I_a^ϵ computed using the private empirical mean of the rewards collected in the last active episode of arm a , Algorithm 1 satisfies ϵ -global DP.*

Proof. Fix two neighboring reward streams $r^T = \{r_1, \dots, r_T\}$ and $r'^T = \{r'_1, \dots, r'_T\}$. This implies that $\exists j \in [1, T]$ such that $r_j \neq r'_j$ and $\forall t \neq j, r_t = r'_t$.

We also fix a sequence of actions $a^T = \{a_1, \dots, a_T\}$.

We want to show that: $Pr(\pi(r^T) = a^T) \leq e^\epsilon Pr(\pi(r'^T) = a^T)$.

The main idea is that the change of reward in the j -th reward only affects the empirical mean computed in one episode, which is made private using the Laplace Mechanism and Lemma 1.

- Since $r^{j-1} = r'^{j-1}$, $Pr(\pi(r^{j-1}) = a^{j-1}) = Pr(\pi(r'^{j-1}) = a^{j-1})$.
- Let $t_\ell \leq j < t_{\ell+1}$ and $t_{\ell'} \leq j < t_{\ell'+1}$ be the episodes corresponding to the j th reward in r^T and r'^T respectively. Since $r^{j-1} = r'^{j-1}$, we get that $\ell = \ell'$. Thus, $Pr(\pi(r^{t_{\ell+1}}) = a^{t_{\ell+1}}) = Pr(\pi(r'^{t_{\ell+1}}) = a^{t_{\ell+1}})$.
- Let $\tilde{\mu}_{a,\epsilon}^\ell$ and $\tilde{\mu}'_{a,\epsilon}^\ell$ be the private means of arm a computed in the episode $[t_\ell, t_{\ell+1}]$, by the Laplace mechanism, for every interval $I \in \mathcal{R}$, $Pr(\tilde{\mu}_{a,\epsilon}^\ell \in I) \leq e^\epsilon Pr(\tilde{\mu}'_{a,\epsilon}^\ell \in I)$.
- Finally, since $\{r_{j+1}, \dots, r_T\} = \{r'_{j+1}, \dots, r'_T\}$, $Pr(\pi(r^T) = a^T | \tilde{\mu}_{a,\epsilon}^\ell \in I) = Pr(\pi(r'^T) = a^T | \tilde{\mu}'_{a,\epsilon}^\ell \in I)$

Now, we conclude the argument by using a chain rule. \square

Since Theorem 6 holds for any index-based bandit algorithm that uses only private empirical means of rewards (Equation (22)) of the last active episode to compute the indices, it also implies that AdaP-UCB and AdaP-KLUCB satisfy ϵ -global DP.

E Upper Bounds on Regret: AdaP-UCB and AdaP-KLUCB

E.1 Concentration Inequalities

Lemma 3. Assume that $(X_i)_{1 \leq i \leq n}$ are iid random variables in $[0, 1]$, with $\mathbb{E}(X_i) = \mu$. Then, for any $\delta \geq 0$,

$$\mathbb{P} \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) - \frac{\log(\frac{1}{\delta})}{n\epsilon} - \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \geq \mu \right) \leq \frac{3}{2}\delta, \quad (23)$$

and

$$\mathbb{P} \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} + \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \leq \mu \right) \leq \frac{3}{2}\delta, \quad (24)$$

where $\hat{\mu}_n = \frac{1}{n} \sum_{t=1}^n X_t$

Proof. We have that

$$\begin{aligned} p_1 &\triangleq \mathbb{P} \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) - \frac{\log(\frac{1}{\delta})}{n\epsilon} - \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \geq \mu \right) \\ &\leq \mathbb{P} \left(\hat{\mu}_n - \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \geq \mu \right) + \mathbb{P} \left(\text{Lap} \left(\frac{1}{n\epsilon} \right) - \frac{\log(\frac{1}{\delta})}{n\epsilon} \geq 0 \right) \\ &\leq \delta + \frac{\delta}{2} = \frac{3}{2}\delta, \end{aligned}$$

where the last inequality is due to Lemma 11 and Lemma 10.

Similarly,

$$\begin{aligned} p_2 &\triangleq \mathbb{P} \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} + \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \leq \mu \right) \\ &\leq \mathbb{P} \left(\hat{\mu}_n + \sqrt{\frac{\log(\frac{1}{\delta})}{2n}} \leq \mu \right) + \mathbb{P} \left(\text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \leq 0 \right) \\ &\leq \delta + \frac{\delta}{2} = \frac{3}{2}\delta, \end{aligned}$$

where the last inequality is due to Lemma 11 and Lemma 10. \square

Lemma 4. Let X_1, X_2, \dots, X_n be a sequence of independent random variables sampled from a Bernoulli distribution with mean μ , and let $\hat{\mu}_n = \frac{1}{n} \sum_{t=1}^n X_t$ be the sample mean. Let

$$\check{\mu}_n(\delta) \triangleq \text{Clip}_{0,1} \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \right) \quad (25)$$

for $\delta > 0$ be the clipped and private empirical mean.

Claim 1. For any $\delta > 0$ and $\alpha \in [0, \mu]$, the following inequality holds:

$$\mathbb{P}(\mu \geq \check{\mu}_n(\delta) + \alpha) \leq \exp(-nd(\mu - \alpha, \mu)) + \frac{1}{2}\delta \quad (26)$$

Claim 2. Furthermore for $\delta \geq 0$, we define

$$U_n(\delta) \triangleq \max \left\{ q \in [0, 1] : d(\check{\mu}_n(\delta), q) \leq \frac{\log(\frac{1}{\delta})}{n} \right\} \quad (27)$$

Then,

$$\mathbb{P}(\mu \geq U_n(\delta)) \leq \frac{3}{2}\delta \quad (28)$$

Proof. Here, we prove Claim 1 followed by Claim 2.

Claim 1. Since $\check{\mu}_n(\delta) = \min \left\{ \max \left\{ 0, \hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \right\}, 1 \right\}$, we have that

$$\begin{aligned} \mu - \alpha \geq \check{\mu}_n(\delta) &\Rightarrow \mu - \alpha \geq 1 \quad \text{or} \quad \mu - \alpha \geq \max \left\{ 0, \left(\hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \right) \right\} \\ &\Rightarrow \mu - \alpha \geq \hat{\mu}_n + \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \quad (\text{since } \mu \leq 1) \\ &\Rightarrow \mu - \alpha \geq \hat{\mu}_n \quad \text{or} \quad \text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \leq 0. \end{aligned}$$

It implies that

$$\begin{aligned} \mathbb{P}(\mu \geq \check{\mu}_n(\delta) + \alpha) &\leq \mathbb{P}(\mu \geq \hat{\mu}_n + \alpha) + \mathbb{P} \left(\text{Lap} \left(\frac{1}{n\epsilon} \right) + \frac{\log(\frac{1}{\delta})}{n\epsilon} \leq 0 \right) \\ &\leq \exp(-nd(\mu - \alpha, \mu)) + \frac{1}{2}\delta. \end{aligned}$$

The last inequality is due to Equation 39 of Lemma 13 and Lemma 10.

Claim 2.

We have that the sets

$$\begin{aligned} \{\mu \geq U_n(\delta)\} &\stackrel{(a)}{=} \{\mu \geq U_n(\delta) \geq \check{\mu}_n(\delta)\} \\ &\stackrel{(b)}{=} \{d(\check{\mu}_n(\delta), \mu) \geq d(\check{\mu}_n(\delta), U_n(\delta)), \mu \geq \check{\mu}_n(\delta)\} \\ &\stackrel{(c)}{=} \{d(\check{\mu}_n(\delta), \mu) \geq \frac{\log(\frac{1}{\delta})}{n}, \mu \geq \check{\mu}_n(\delta)\} \\ &\stackrel{(d)}{=} \{\check{\mu}_n(\delta) \leq \mu - \alpha\} \end{aligned}$$

Here, we chose an $\alpha > 0$ such that $d(\mu - \alpha, \mu) = \frac{\log(\frac{1}{\delta})}{n}$.

Step (a) holds because $U_n(\delta) \geq \check{\mu}_n(\delta)$ by the definition of $U_n(\delta)$. Step (b) also holds true since $d(\check{\mu}_n(\delta), \cdot)$ is increasing on $[\check{\mu}_n(\delta), 1]$. Since $d(\check{\mu}_n(\delta), U_n(\delta)) = \frac{\log(\frac{1}{\delta})}{n}$ by the definition of $U_n(\delta)$, we obtain the equality in Step (c). Finally, Step (d) is obtained by inverting the relative entropy.

We conclude the proof by

$$\begin{aligned} \mathbb{P}\{\mu \geq U_n(\delta)\} &= \mathbb{P}\{\check{\mu}_n(\delta) \leq \mu - \alpha\} \\ &\leq \exp(-nd(\mu - \alpha, \mu)) + \frac{1}{2}\delta \quad (\text{by Claim 1}) \\ &= \delta + \frac{\delta}{2} = \frac{3}{2}\delta \quad (\text{by substituting } \alpha) \end{aligned}$$

□

E.2 Generic Regret Analysis for Algorithm 1

Algorithm 1 is a generic framework to construct an extension of any optimistic index-based bandit algorithm, which would satisfy ϵ -global DP. The algorithm is based on the index I_a^ϵ of each arm. I_a^ϵ is computed using the private empirical mean of the last active episode of arm a and is a high probability upper bound of the real mean μ_a .

To explicate the two conditions on arm indexes, we introduce the notation $I_a^\epsilon(t-1, \alpha, s)$, which is the index of arm a , at time-step t and computed using s reward samples from arm a .

Thus, we can express the index computed using just the last active episode as

$$I_a^\epsilon(t-1, \alpha) = I_a^\epsilon(t-1, \alpha, \frac{1}{2}N_a(t-1)). \quad (29)$$

Because $I_a^\epsilon(t-1, \alpha)$ only uses samples collected from the last active episode, and due to the doubling, the last active episode's size is exactly half the number of times arm a was pulled since the beginning.

The optimism of the index is ensured by the fact that

$$\mathbb{P}(I_a^\epsilon(t-1, \alpha, s) \leq \mu_a) \leq \frac{3}{2} \frac{1}{t^\alpha} \quad (30)$$

for every arm a , every sample size s and every time-step t , where α is the confidence level.

Theorem 11. *Let a be a suboptimal arm and $\ell \in \mathbb{N}$ such that $2^\ell < T$. Then, Algorithm 1 using an index I_a^ϵ satisfying Equations 29 and 30, also satisfies that for any $\alpha > 3$,*

$$\mathbb{E}[N_a(T)] \leq 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T + \frac{\alpha}{\alpha - 3},$$

where $G_{a,\ell,T} = \{I_a^\epsilon(T-1, \alpha, 2^\ell) < \mu^*\}$ and $G_{a,\ell,T}^c$ is the complement of $G_{a,\ell,T}$

Proof. Without loss of generality, we assume the first arm is the optimal one ($\mu^* = \mu_1$) and denote a suboptimal arm by a ($1 < a \leq K$).

We leverage the standard idea of UCB-type proofs: if arm a is chosen at the beginning of an episode ℓ , then either its index at t_ℓ is larger than the true mean of the first arm, or the true mean of the first arm is larger than the first arm's index at t_ℓ .

Since decisions, i.e. playing the arm with the highest index, are only taken at the beginning of an episode, we introduce ϕ which takes as input a time step and outputs the time step corresponding to the beginning of an episode. Formally, for each $t \in [K+1, T]$, let $\phi(t) = t_\ell$ such that $t_\ell \leq t \leq t_{\ell+1} - 1$. In Example 2, $\phi(5) = 4$ and $\phi(9) = 7$.

Formally, $\phi(t)$ is a random variable such that

$$\forall t : \phi(t) \leq t \leq 2\phi(t) \quad (31)$$

Step 1: Decomposition of $N_a(T)$. We observe that

$$\begin{aligned} N_a(T) &= 1 + \sum_{t=K+1}^T \mathbb{I}\{A_t = a\} \\ &= 1 + \sum_{t=K+1}^T \mathbb{I}\{A_t = a \text{ and } I_1^\epsilon(\phi(t) - 1, \alpha) > \mu_1\} + \mathbb{I}\{A_t = a \text{ and } I_1^\epsilon(\phi(t) - 1, \alpha) \leq \mu_1\} \\ &\leq 1 + \underbrace{N'_a(T)}_{\text{Term1}} + \underbrace{\sum_{t=K+1}^T \mathbb{I}\{I_1^\epsilon(\phi(t) - 1, \alpha) \leq \mu_1\}}_{\text{Term2}} \end{aligned}$$

We define $N'_a(T) \triangleq \sum_{t=K+1}^T \mathbb{I}\{A_t = a \text{ and } I_1^\epsilon(t_{\ell'} - 1, \alpha) > \mu_1\}$

Step 2: Decomposition of Term 1: $N'_a(T)$. Let $G_{a,\ell,T}$ be the 'good' event defined by

$$G_{a,\ell,T} = \{I_a^\epsilon(T-1, \alpha, 2^\ell) < \mu_1\}.$$

The main part of the proof is decomposing $N'_a(T)$ among the 'good' and the 'bad' events, i.e.

$$\mathbb{E}[N'_a(T)] = \mathbb{E}[\mathbb{I}\{G_{a,\ell,T}\} N'_a(T)] + \mathbb{E}[\mathbb{I}\{G_{a,\ell,T}^c\} N'_a(T)] \leq 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T.$$

$G_{a,\ell,T}^c$ denotes the complement of $G_{a,\ell,T}$.

To prove the last inequality, we only need to prove that when $G_{a,\ell,T}$ happens, $N'_a(T) \leq 2^{\ell+1}$. We prove it by contradiction.

Hence, let us assume that $G_{a,\ell,T}$ holds but $N'_a(T) > 2^{\ell+1}$.

This assumption implies that the arm a is played more than $2^{\ell+1}$ times. Thus, there must exist a round $t_{\ell'}$, where $N_a(t_{\ell'} - 1) = 2^{\ell+1}$, $A_{t_{\ell'}} = a$ and $I_1^\epsilon(t_{\ell'} - 1, \alpha) \geq \mu_1$. Since indices are computed only

using the samples from the last active episode, $I_a^\epsilon(t_{\ell'} - 1, \alpha)$ is computed using exactly 2^ℓ reward samples from arm a .

Thus, we obtain

$$\begin{aligned} I_a^\epsilon(t_{\ell'} - 1, \alpha) &= I_a^\epsilon(t_{\ell'} - 1, \alpha, 2^\ell) \\ &\leq I_a^\epsilon(T - 1, \alpha, 2^\ell) \quad (\text{because } t_{\ell'} \leq T \text{ and } I_a^\epsilon(\cdot, \alpha, 2^\ell) \text{ is increasing}) \\ &< \mu_1 \quad (\text{definition of } G_{a,\ell,T}) \\ &\leq I_1^\epsilon(t_{\ell'} - 1, \alpha) \end{aligned}$$

The last inequality contradicts the fact that $A_{t_{\ell'}} = i$ and thus, establishes the claim that $N_a'(T) \leq 2^{\ell+1}$ under the ‘good’ event.

Step 3: Upper-bounding Term 2. To conclude,

$$\begin{aligned} \mathbb{E} \left[\sum_{t=K+1}^T \mathbb{I}\{I_1^\epsilon(\phi(t) - 1, \alpha) \leq \mu_1\} \right] &= \sum_{t=K+1}^T \mathbb{P}\{I_1^\epsilon(\phi(t) - 1, \alpha) \leq \mu_1\} \\ &\leq \sum_{t=K+1}^T \sum_{\phi=t/2}^t \mathbb{P}\{I_1^\epsilon(\phi - 1, \alpha) \leq \mu_1\} \\ &\leq \sum_{t=K+1}^T \sum_{\phi=t/2}^t \sum_{s=1}^{\phi} \mathbb{P}\{I_1^\epsilon(\phi - 1, \alpha, s) \leq \mu_1\} \\ &\leq \sum_{t=K+1}^T \sum_{\phi=t/2}^t \sum_{s=1}^{\phi} \frac{3}{2} \frac{1}{\phi^\alpha} \quad (\text{Equation 30}) \\ &= \frac{3}{2} \sum_{t=K+1}^T \sum_{\phi=t/2}^t \frac{1}{\phi^{\alpha-1}} \\ &\leq \frac{3}{2} \sum_{t=K+1}^T \frac{2^{\alpha-2}}{t^{\alpha-2}} \quad (\text{because } \phi \geq \frac{t}{2}) \\ &\leq \frac{3}{2} 2^{\alpha-2} \int_K^T \frac{1}{x^{\alpha-2}} dx \quad (\text{sum-integral inequality}) \\ &\leq \frac{3}{2} 2^{\alpha-2} \frac{1}{\alpha-3} \frac{1}{K^{\alpha-3}} = \frac{3}{2} \frac{2}{\alpha-3} \left(\frac{2}{K}\right)^{\alpha-3} \\ &\leq \frac{3}{\alpha-3} \end{aligned}$$

for $\alpha > 3$ and $K \geq 2$.

Here, the first inequality is due to an union bound on $\phi(t) \in [t/2, t]$ (Equation 31), and the second inequality is due to a union bound on $N_1(\phi - 1)$.

Step 4: Combining the Bounds on Terms 1 and 2.

$$\begin{aligned} \mathbb{E}[N_a(T)] &\leq 1 + 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T + \frac{3}{\alpha-3} \\ &= 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T + \frac{\alpha}{\alpha-3} \end{aligned}$$

□

Now we design indexes that satisfy the conditions of Theorem 11, namely AdaP-UCB and AdaP-KLUCB.

To obtain the final regret bounds, we only have to choose ℓ big enough such that $\mathbb{P}(I_a(T, 2^\ell) \geq \mu_1) T$ is negligible. This corresponds to the leading term in the regret upper-bounds, and this is where the regrets of AdaP-UCB and AdaP-KLUCB differ.

We explicate the issues of designing the indexes and choosing corresponding ℓ in the following section, which leads to the regret upper bounds of AdaP-UCB and AdaP-KLUCB.

E.3 Regret Analysis for AdaP-UCB and AdaP-KLUCB

Theorem 7. *For rewards in $[0, 1]$, AdaP-UCB satisfies ϵ -global DP, and for $\alpha > 3$, it yields a regret*

$$\text{Reg}_T(\text{AdaP-UCB}, \nu) \leq \sum_{a: \Delta_a > 0} \left(\frac{16\alpha}{\min\{\Delta_a, \epsilon\}} \log(T) + \frac{3\alpha}{\alpha - 3} \right).$$

Proof. The proof is constituted of three steps.

Step 1: Designing an Index satisfying Equation (29), Equation (30), and ϵ -global DP. For AdaP-UCB, the index is defined as

$$I_a^\epsilon(t_\ell - 1, \alpha) = \tilde{\mu}_{a,\epsilon}^\ell + \sqrt{\frac{\alpha \log(t_\ell)}{2 \times \frac{1}{2} N_a(t_\ell - 1)}} + \frac{\alpha \log(t_\ell)}{\epsilon \times \frac{1}{2} N_a(t_\ell - 1)},$$

where

$$\tilde{\mu}_{a,\epsilon}^\ell = \hat{\mu}_{a, \frac{1}{2} N_a(t_\ell - 1)} + \text{Lap} \left(\frac{1}{\epsilon \times \frac{1}{2} N_a(t_\ell - 1)} \right) \quad (32)$$

is the private empirical mean of arm a computed using only samples from the last active episode, and $\hat{\mu}_{a,s}$ is the empirical mean of arm a calculated using s samples of reward from arm a .

This index verifies the first condition (Equation 29) of Theorem 11.

The second condition (Equation 30) of Theorem 11 follows directly from Equation 24 of Lemma 3

By Theorem 6, AdaP-UCB is ϵ -global DP.

By Theorem 11, for every suboptimal arm a , we have that

$$\mathbb{E}[N_a(T)] \leq 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T + \frac{\alpha}{\alpha - 3},$$

where

$$G_{a,\ell,T} = \left\{ \hat{\mu}_{a,2^\ell} + \text{Lap} \left(\frac{1}{2^\ell \epsilon} \right) + \sqrt{\frac{\alpha \log(T)}{2 \times 2^\ell}} + \frac{\alpha \log(T)}{\epsilon 2^\ell} < \mu_1 \right\}.$$

Step 2: Choosing an ℓ . Now, we observe that

$$\begin{aligned} \mathbb{P}(G_{a,\ell,T}^c) &= \mathbb{P} \left(\hat{\mu}_{a,2^\ell} + \text{Lap} \left(\frac{1}{2^\ell \epsilon} \right) + \sqrt{\frac{\alpha \log(T)}{2 \times 2^\ell}} + \frac{\alpha \log(T)}{\epsilon 2^\ell} \geq \mu_1 \right) \\ &= \mathbb{P} \left(\hat{\mu}_{a,2^\ell} + \text{Lap} \left(\frac{1}{2^\ell \epsilon} \right) - \sqrt{\frac{\alpha \log(T)}{2 \times 2^\ell}} - \frac{\alpha \log(T)}{\epsilon 2^\ell} \geq \mu_a + \gamma \right) \end{aligned}$$

for $\gamma = \left(\Delta_a - 2\sqrt{\frac{\alpha \log(T)}{2 \times 2^\ell}} - 2\frac{\alpha \log(T)}{\epsilon 2^\ell} \right)$.

The idea is to choose ℓ big enough so that $\gamma \geq 0$.

Let us consider the contrary, i.e.

$$\begin{aligned} \gamma < 0 &\Rightarrow \sqrt{2^\ell} < \sqrt{\frac{\alpha \log(T)}{2\Delta_a^2}} \left(1 + \sqrt{1 + \frac{4\Delta_a}{\epsilon}} \right) \\ &\Rightarrow 2^\ell < \frac{\alpha \log(T)}{2\Delta_a^2} \left(4 + \frac{8\Delta_a}{\epsilon} \right) \\ &\Rightarrow 2^\ell < \frac{4\alpha \log(T)}{\Delta_a \min\{\epsilon, 2\Delta_a\}}. \end{aligned} \quad (33)$$

Thus, by choosing

$$\ell = \left\lceil \frac{1}{\log(2)} \log \left(\frac{4\alpha \log(T)}{\Delta_a \min\{\epsilon, 2\Delta_a\}} \right) \right\rceil$$

we ensure $\gamma > 0$. This also implies that

$$\mathbb{P}(G_{a,\ell,T}^c) \leq \mathbb{P} \left(\hat{\mu}_{a,2^\ell} + \text{Lap} \left(\frac{1}{2^\ell \epsilon} \right) - \sqrt{\frac{\alpha \log(T)}{2 \times 2^\ell}} - \frac{\alpha \log(T)}{\epsilon 2^\ell} \geq \mu_a \right) \leq \frac{3}{2T^\alpha}$$

The last inequality is due to Equation 23 of Lemma 3.

Step 3: The Regret Bound. Combining Steps 1 and 2, we get that

$$\begin{aligned} \mathbb{E}[N_a(T)] &\leq \frac{\alpha}{\alpha - 3} + 2^{\ell+1} + T \times \frac{3}{2T^\alpha} \\ &\leq \frac{16\alpha \log(T)}{\Delta_a \min\{\epsilon, 2\Delta_a\}} + \frac{3\alpha}{\alpha - 3}. \end{aligned} \quad (34)$$

Plugging this upper bound back in the definition of problem-dependent regret concludes the proof. \square

Remark 3. The leading term of the regret is $\frac{16\alpha \log(T)}{\Delta_a \min\{\epsilon, 2\Delta_a\}}$, which is 4 times more than what we got from Equation 33. A multiplicative factor of 2 is introduced due to the doubling and another multiplicative factor of 2 is due to the forgetting. Thus, the combined price of doubling and forgetting is a multiplicative constant 4 in the leading term of regret.

Theorem 8. When the rewards are sampled from Bernoulli distributions, AdaP-KLUCB satisfies ϵ -global DP, and for $\alpha > 3$ and constants $C_1(\alpha), C_2 > 0$, it yields a regret

$$\text{Reg}_T(\text{AdaP-KLUCB}, \nu) \leq \sum_{a: \Delta_a > 0} \left(\frac{C_1(\alpha) \Delta_a}{\min\{d(\mu_a, \mu^*), C_2 \epsilon \Delta_a\}} \log(T) + \frac{\alpha}{\alpha - 3} \right).$$

Proof. The proof is constituted of three steps.

Step 1: Designing an Index satisfying Equation (29), Equation (30), and ϵ -global DP. For AdaP-KLUCB, the index is defined as

$$I_a^\epsilon(t_\ell - 1, \alpha) = \max \left\{ q \in [0, 1] : d(\check{\mu}_{a,\epsilon}^\ell, q) \leq \frac{\alpha \log(t_\ell)}{\frac{1}{2} N_a(t_\ell - 1)} \right\} \triangleq U_{a, \frac{1}{2} N_a(t_\ell - 1)} \left(\frac{1}{t_\ell^\alpha} \right),$$

where $\check{\mu}_{a,\epsilon}^\ell = \text{Clip}_{0,1} \left(\tilde{\mu}_{a,\epsilon}^\ell + \frac{\alpha \log(t_\ell)}{\epsilon \frac{1}{2} N_a(t_\ell - 1)} \right) = \check{\mu}_{a, \frac{1}{2} N_a(t_\ell - 1)} \left(\frac{1}{t_\ell^\alpha} \right)$ as defined in Equation 25,

$\tilde{\mu}_{a,\epsilon}^\ell$ is the private empirical computed only using the samples from the last active episode (as defined for AdaP-UCB, and $U_{a,s}(\delta) = \max \left\{ q \in [0, 1] : d(\check{\mu}_{a,s}(\delta), q) \leq \frac{\log(\frac{1}{s})}{s} \right\}$ as defined in Equation 27

This index verifies the first condition (Equation 29) of Theorem 11.

The second condition (Equation 30) of Theorem 11 follows directly from Equation 24 of Lemma 3 By Theorem 6, AdaP-KLUCB also satisfies ϵ -global DP.

By Theorem 11, for every suboptimal arm a , we have that

$$\mathbb{E}[N_a(T)] \leq 2^{\ell+1} + \mathbb{P}(G_{a,\ell,T}^c) T + \frac{\alpha}{\alpha - 3},$$

where

$$G_{a,\ell,T} = \left\{ U_{a,2^\ell} \left(\frac{1}{T^\alpha} \right) < \mu_1 \right\}.$$

Step 2: Choosing an ℓ . We observe that

$$\begin{aligned}\mathbb{P}(G_{a,\ell,T}^c) &= \mathbb{P}\left(U_{a,2^\ell} \left(\frac{1}{T^\alpha}\right) \geq \mu_1\right) \\ &\leq \mathbb{P}\left(d^+ \left(\check{\mu}_{a,2^\ell} \left(\frac{1}{T^\alpha}\right), \mu_1\right) \leq \frac{\alpha \log(T)}{2^\ell}\right) \quad (\text{by definition of } U_{a,2^\ell})\end{aligned}$$

where $d^+(p, q) \triangleq d(p, q) \mathbb{I}_{p < q}$ and $d(p, q)$ is the relative entropy between Bernoulli distributions as stated in Definition 5.

Let $\beta > 0$, and $c(\beta) \in [0, 1]$ such that: $d(\mu_a + c(\beta)\Delta_a, \mu_1) = \frac{d(\mu_a, \mu_1)}{1+\beta}$.

Since $d(\cdot, \mu_1)$ is a bijective function from $[\mu_a, \mu_1]$ to $[0, d(\mu_a, \mu_1)]$, we get that $c(\beta)$ always exists and is unique.

In addition, $c(\beta)$ verifies: $\lim_{\beta \rightarrow 0} c(\beta) = 0$, $\lim_{\beta \rightarrow +\infty} c(\beta) = 1$ and $c(\beta)$ is an increasing function of β .

First, we choose ℓ such that

$$2^\ell \geq \frac{(1+\beta)\alpha \log(T)}{d(\mu_a, \mu_1)}. \quad (35)$$

This leads to

$$\begin{aligned}\mathbb{P}(G_{a,\ell,T}^c) &\leq \mathbb{P}\left(d^+ \left(\check{\mu}_{a,2^\ell} \left(\frac{1}{T^\alpha}\right), \mu_1\right) \leq \frac{d(\mu_a, \mu_1)}{1+\beta}\right) \\ &= \mathbb{P}\left(d^+ \left(\check{\mu}_{a,2^\ell} \left(\frac{1}{T^\alpha}\right), \mu_1\right) \leq d(\mu_a + c(\beta)\Delta_a, \mu_1)\right) \quad (\text{definition of } c(\beta)) \\ &\leq \mathbb{P}\left(\check{\mu}_{a,2^\ell} \left(\frac{1}{T^\alpha}\right) \geq \mu_a + c(\beta)\Delta_a\right) \quad (d(\cdot, \mu_1) \text{ is decreasing on } [0, \mu_1]) \\ &\leq \mathbb{P}\left(\hat{\mu}_{a,2^\ell} + \text{Lap}\left(\frac{1}{2^\ell \epsilon}\right) + \frac{\alpha \log(T)}{\epsilon 2^\ell} \geq \mu_a + c(\beta)\Delta_a\right) \quad (\text{definition of } \check{\mu})\end{aligned}$$

Let us consider $\gamma_{\ell,T}$ such that $d(\mu_a + \gamma_{\ell,T}\Delta_a, \mu_a) = \frac{\log(T)}{2^\ell}$. We prove its existence and upper bound it later in Fact 3. Thus, we obtain

$$\begin{aligned}\mathbb{P}(G_{a,\ell,T}^c) &\leq \mathbb{P}\left(\hat{\mu}_{a,2^\ell} - \gamma_{\ell,T}\Delta_a + \text{Lap}\left(\frac{1}{2^\ell \epsilon}\right) - \frac{\log(T)}{\epsilon 2^\ell} \geq \mu_a + (c(\beta) - \gamma_{\ell,T})\Delta_a - \frac{(1+\alpha)\log(T)}{\epsilon 2^\ell}\right) \\ &= \mathbb{P}\left(\hat{\mu}_{a,2^\ell} - \gamma_{\ell,T}\Delta_a + \text{Lap}\left(\frac{1}{2^\ell \epsilon}\right) - \frac{\log(T)}{\epsilon 2^\ell} \geq \mu_a + \theta\right)\end{aligned}$$

Here, $\theta \triangleq (c(\beta) - \gamma_{\ell,T})\Delta_a - \frac{(1+\alpha)\log(T)}{\epsilon 2^\ell}$.

By choosing

$$2^\ell \geq \frac{(1+\alpha)\log(T)}{(c(\beta) - \gamma_{\ell,T})\epsilon \Delta_a}, \quad (36)$$

we ensure that $\theta \geq 0$. Thus, we get

$$\begin{aligned}\mathbb{P}(G_{a,\ell,T}^c) &\leq \mathbb{P}\left(\hat{\mu}_{a,2^\ell} - \gamma_{\ell,T}\Delta_a + \text{Lap}\left(\frac{1}{2^\ell \epsilon}\right) - \frac{\log(T)}{\epsilon 2^\ell} \geq \mu_a\right) \\ &\leq \mathbb{P}\left(\hat{\mu}_{a,2^\ell} - \gamma_{\ell,T}\Delta_a \geq \mu_a\right) + \mathbb{P}\left(\text{Lap}\left(\frac{1}{2^\ell \epsilon}\right) - \frac{\log(T)}{\epsilon 2^\ell} \geq 0\right) \\ &\leq \exp(-2^\ell d(\mu_a + \gamma_{\ell,T}\Delta_a, \mu_a)) + \frac{1}{2T} \\ &= \frac{3}{2T}.\end{aligned}$$

The last inequality is due to Equation 38 of Lemma 13 and Lemma 10.

Fact 3. B $\triangleq \{\beta > 0 : c(\beta) > \gamma_{\ell,T}\} \neq \emptyset$.

Combining both conditions 35 and 35, we choose ℓ to be the smallest integer such that

$$2^\ell \geq \inf_{\beta \in \mathbf{B}} \max \left\{ \frac{(1+\beta)\alpha}{d(\mu_a, \mu_1)}, \frac{(1+\alpha)}{(c(\beta) - \gamma_{\ell, T})\epsilon\Delta_a} \right\} \log(T) \triangleq \frac{\frac{1}{4}C_1(\alpha)}{\min\{d(\mu_a, \mu_1), C_2\epsilon\Delta_a\}} \log(T)$$

Step 3: The Regret Bound. Combining Steps 1 and 2, we get that

$$\begin{aligned} \mathbb{E}[N_a(T)] &\leq 2^{\ell+1} + T \times \frac{3}{2T} + \frac{\alpha}{\alpha-3} \\ &\leq \frac{C_1(\alpha)}{\min\{d(\mu_a, \mu_1), C_2\epsilon\Delta_a\}} \log(T) + \frac{3\alpha}{\alpha-3} \end{aligned}$$

Plugging this upper bound back in the definition of problem-dependent regret concludes the proof. \square

To conclude, we prove Fact 3.

Fact 3. $\mathbf{B} \triangleq \{\beta > 0 : c(\beta) > \gamma_{\ell, T}\} \neq \emptyset$.

Proof. Step 1: Going from $d(\cdot, \mu_a)$ to $d(\cdot, \mu_1)$. The difficulty of the proof lies in the fact that $\gamma_{\ell, T}$ is defined by inverting $d(\cdot, \mu_a)$ while $c(\beta)$ is defined by inverting $d(\cdot, \mu_1)$.

To handle that, we investigate the function $g(x) \triangleq d(x, \mu_a) - d(x, \mu_1)$.

g satisfies the following properties:

- g is continuous and increasing in the interval $[\mu_a, \mu_1]$,
- $g(\mu_a) = -d(\mu_a, \mu_1) < 0$, and
- $g(\mu_1) = d(\mu_1, \mu_a) > 0$.

This implies that there exists a unique root of $g(x)$, where it changes sign. Specifically, there exists a unique $z \in [\mu_a, \mu_1]$ such that:

- $g(z) = 0$
- $\forall x \in [\mu_a, z[: g(x) < 0$
- $\forall x \in]z, \mu_1] : g(x) > 0$

and consequently z verifies $d(z, \mu_a) = d(z, \mu_1)$

Step 2: Choosing β . We choose β such that $\frac{d(\mu_a, \mu_1)}{1+\beta} = d(z, \mu_a) = d(z, \mu_1)$.

Step 3: Consequence of the choice of β on $c(\beta)$. Thus,

$$d(\mu_a + c(\beta)\Delta_a, \mu_1) = d(z, \mu_1),$$

which yields

$$z = \mu_a + c(\beta)\Delta_a$$

by uniqueness of z .

Step 4: Consequence of the choice of β on $\gamma_{\ell, T}$. On the other hand,

$$\begin{aligned} d(\mu_a + \gamma_{\ell, T}\Delta_a, \mu_a) &= \frac{\log(T)}{2^\ell} && \text{(by definition of } \gamma_{\ell, T}\text{)} \\ &\leq \frac{d(\mu_a, \mu_1)}{\alpha(\beta+1)} && \text{(by Equation 35)} \\ &< d(z, \mu_a) && \text{(since } \alpha > 3\text{)} \\ &= d(\mu_a + c(\beta)\Delta_a, \mu_a) && \text{(37)} \end{aligned}$$

As a consequence, we conclude that $\gamma_{\ell, T}$ exists and $\gamma_{\ell, T} < c(\beta)$ as $d(\cdot, \mu_a)$ is an increasing function in the interval $[\mu_a, 1]$ \square

E.4 Problem-independent Regret Bounds

In this section, we provide problem-independent (or minimax) regret upper bounds for AdaP-UCB.

Theorem 12. *For rewards in $[0, 1]$, AdaP-UCB yields a regret*

$$\text{Reg}_T(\text{AdaP-UCB}, \nu) \leq \frac{3\alpha}{\alpha-3} \sum_a \Delta_a + 8\sqrt{\alpha K T \log(T)} + \frac{16\alpha K \log(T)}{\epsilon}$$

which achieves the minimax lower bound of Thm 2 up to $\log(T)$ factors.

Proof. Let Δ be a value to be tuned later.

We have that

$$\begin{aligned} \text{Reg}_T(\text{AdaP-UCB}, \nu) &= \sum_a \Delta_a \mathbb{E}[N_a(T)] = \sum_{a:\Delta_a \leq \Delta} \Delta_a \mathbb{E}[N_a(T)] + \sum_{a:\Delta_a > \Delta} \Delta_a \mathbb{E}[N_a(T)] \\ &\leq T\Delta + \sum_{a:\Delta_a > \Delta} \Delta_a \left(\frac{16\alpha \log(T)}{\Delta_a \min\{\epsilon, \Delta_a\}} + \frac{3\alpha}{\alpha-3} \right) \quad (\text{eq. 34}) \\ &\leq T\Delta + \frac{16\alpha K \log(T)}{\Delta} + \frac{16\alpha K \log(T)}{\epsilon} + \frac{3\alpha}{\alpha-3} \sum_a \Delta_a \\ &\leq 8\sqrt{\alpha K T \log(T)} + \frac{16\alpha K \log(T)}{\epsilon} + \frac{3\alpha}{\alpha-3} \sum_a \Delta_a \end{aligned}$$

where the last step is by taking $\Delta = 4\sqrt{\frac{\alpha K \log(T)}{T}}$.

□

Remark 4. *The same bound is achieved by AdaP-KLUCB (up to multiplicative constants) by using that $d(\mu_a, \mu^*) \geq 2\Delta_a^2$ and using the same steps in Thm 12.*

F Existing Technical Results and Definitions

In this section, we summarise the existing technical results and definitions required to establish our proofs.

Lemma 5 (Post-processing Lemma (Proposition 2.1, (Dwork and Roth, 2014))). *If a randomised algorithm \mathcal{A} satisfies (ϵ, δ) -Differential Privacy and f is an arbitrary randomised mapping defined on \mathcal{A} 's output, then $f \circ \mathcal{A}$ satisfies (ϵ, δ) -DP.*

Lemma 6 (Markov's Inequality). *For any random variable X and $\epsilon > 0$,*

$$\mathbb{P}(|X| \geq \epsilon) \leq \frac{\mathbb{E}[|X|]}{\epsilon}.$$

Definition 4 (Consistent Policies). *A policy π is called consistent over a class of bandits \mathcal{E} if for all $\nu \in \mathcal{E}$ and $p > 0$, it holds that*

$$\lim_{T \rightarrow \infty} \frac{\text{Reg}_T(\pi, \nu)}{T^p} = 0.$$

The class of consistent policies over \mathcal{E} is denoted by $\Pi_{\text{cons}}(\mathcal{E})$.

Lemma 7 (Divergence decomposition). *Let $\nu = (P_1, \dots, P_K)$ and $\nu' = (P'_1, \dots, P'_K)$ be two bandit instances. Fix some policy π and let $\mathbb{P}_{\nu\pi}$ and $\mathbb{P}_{\nu'\pi}$ be the probability measures on the canonical bandit model. Then,*

$$D_{\text{KL}}(\mathbb{P}_{\nu\pi} \parallel \mathbb{P}_{\nu'\pi}) = \sum_{a=1}^K \mathbb{E}_{\nu} [N_a(T)] D(P_a, P'_a).$$

Lemma 8 (Bretagnolle-Huber inequality). *Let \mathbb{P} and \mathbb{Q} be probability measures on the same measurable space (Ω, \mathcal{F}) , and let $A \in \mathcal{F}$ be an arbitrary event. Then,*

$$\mathbb{P}(A) + \mathbb{Q}(A^c) \geq \frac{1}{2} \exp(-D(\mathbb{P}, \mathbb{Q})),$$

where $A^c = \Omega \setminus A$ is the complement of A .

Lemma 9 (Pinsker's Inequality). *For two probability measures \mathbb{P} and \mathbb{Q} on the same probability space (Ω, \mathcal{F}) , we have*

$$D_{\text{KL}}(\mathbb{P} \parallel \mathbb{Q}) \geq 2(\text{TV}(\mathbb{P} \parallel \mathbb{Q}))^2.$$

Lemma 10 (Tail Bounds for Laplacian Random Variables). *For any $a, b > 0$, we have*

$$\mathbb{P}(\text{Lap}(b) > a) = \frac{1}{2} \exp\left(-\frac{a}{b}\right) \quad \text{and} \quad \mathbb{P}(\text{Lap}(b) < -a) = \frac{1}{2} \exp\left(-\frac{a}{b}\right).$$

Lemma 11 (Hoeffding's Bound). *Assume that $(X_i)_{1 \leq i \leq n}$ are iid random variables in $[0, 1]$, with $\mathbb{E}(X_i) = \mu$. For any $\delta, \beta \geq 0$ and, we have:*

$$\mathbb{P}(\hat{\mu}_n \geq \mu + \beta) \leq \exp(-2n\beta^2) \quad \text{and} \quad \mathbb{P}(\hat{\mu}_n \leq \mu - \beta) \leq \exp(-2n\beta^2),$$

where $\hat{\mu}_n = \frac{1}{n} \sum_{t=1}^n X_t$.

Definition 5 (Relative entropy between Bernoulli distributions). *The relative entropy between Bernoulli distributions with parameters $p, q \in [0, 1]$ is*

$$d(p, q) = p \log(p/q) + (1-p) \log((1-p)/(1-q)),$$

where singularities are defined by taking limits: $d(0, q) = \log(1/(1-q))$ and $d(1, q) = \log(1/q)$ for $q \in [0, 1]$ and $d(p, 0) = 0$ if $p = 0$ and ∞ otherwise and $d(p, 1) = 0$ if $p = 1$ and ∞ otherwise.

Lemma 12 (Properties of the relative entropy between Bernoulli distributions (Lemma 10.2, (Lattimore and Szepesvári, 2018))). *Let $p, q, \epsilon \in [0, 1]$.*

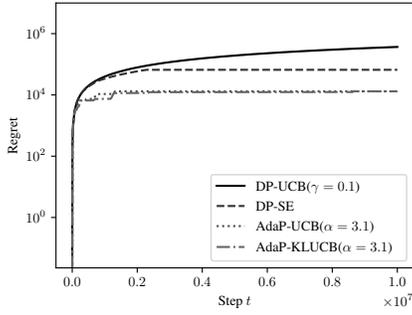
1. *The functions $d(\cdot, q)$ and $d(p, \cdot)$ are convex and have unique minimisers at q and p , respectively.*
2. *$d(p, \cdot)$ and $d(\cdot, p)$ are increasing in the interval $[p, 1]$ and decreasing in the interval $[0, p]$.*

Lemma 13 (Chernoff's Bound). *Let X_1, X_2, \dots, X_n be a sequence of independent random variables that are Bernoulli distributed with mean μ , and let $\hat{\mu}_n = \frac{1}{n} \sum_{t=1}^n X_t$ be the sample mean. Then, for $\beta \in [0, 1 - \mu]$, it holds that:*

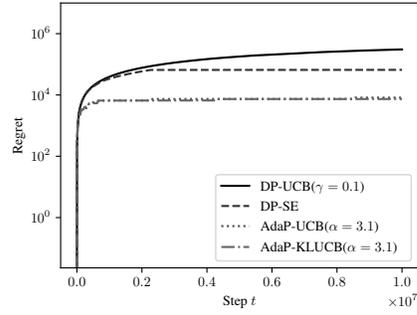
$$\mathbb{P}(\hat{\mu}_n \geq \mu + \beta) \leq \exp(-nd(\mu + \beta, \mu)), \tag{38}$$

and for $\beta \in [0, \mu]$,

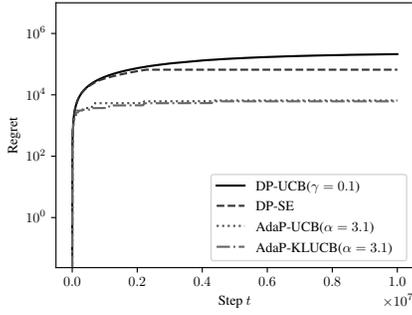
$$\mathbb{P}(\hat{\mu}_n \leq \mu - \beta) \leq \exp(-nd(\mu - \beta, \mu)). \tag{39}$$



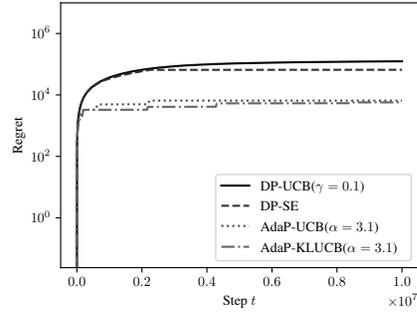
(a) $\epsilon = 0.1$



(b) $\epsilon = 0.25$

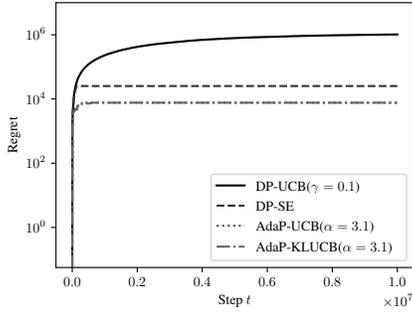


(c) $\epsilon = 0.5$

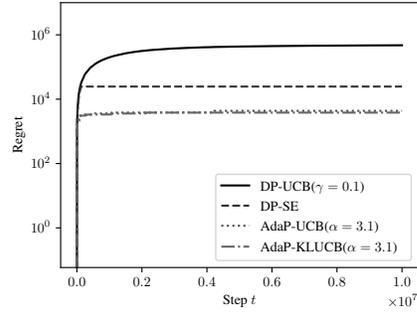


(d) $\epsilon = 1$

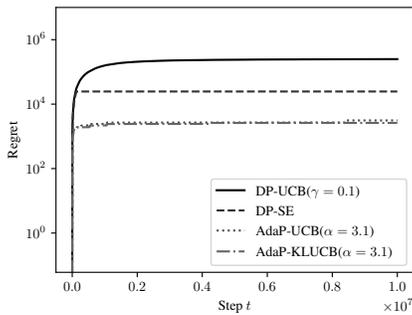
Figure 4: Evolution of regret over time for DP-UCB, DP-SE, AdaP-UCB, and AdaP-KLUCB under C_1 for different values of the privacy budget ϵ . AdaP-KLUCB achieves the lowest regret.



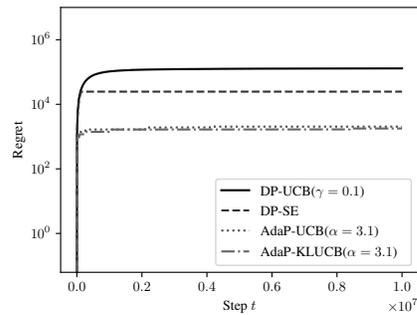
(a) $\epsilon = 0.1$



(b) $\epsilon = 0.25$

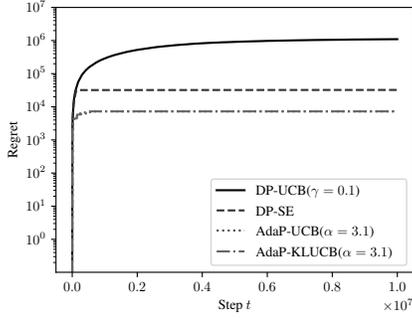


(c) $\epsilon = 0.5$

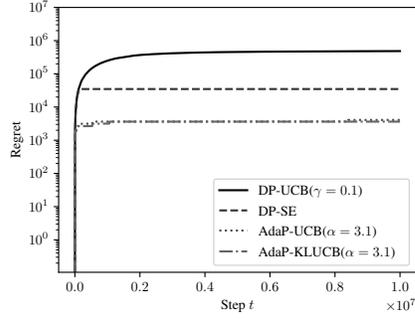


(d) $\epsilon = 1$

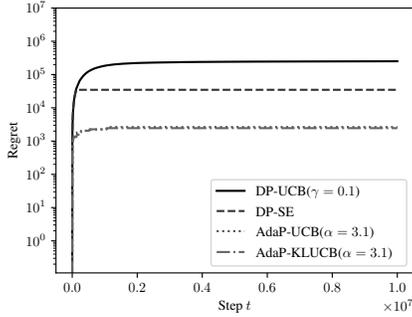
Figure 5: Evolution of regret over time for DP-UCB, DP-SE, AdaP-UCB, and AdaP-KLUCB under C_2 for different values of the privacy budget ϵ . AdaP-KLUCB achieves the lowest regret.



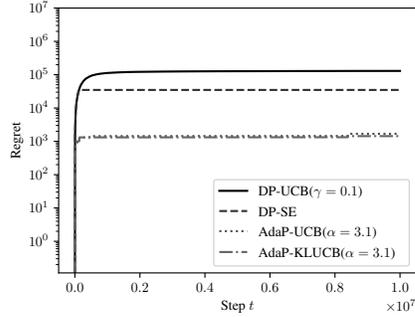
(a) $\epsilon = 0.1$



(b) $\epsilon = 0.25$

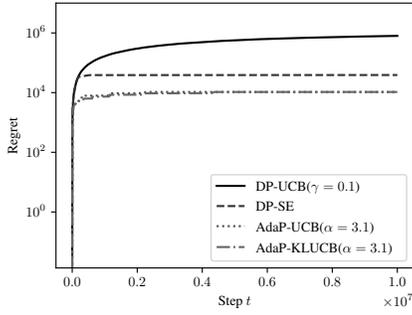


(c) $\epsilon = 0.5$

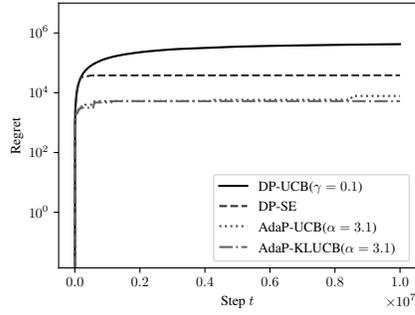


(d) $\epsilon = 1$

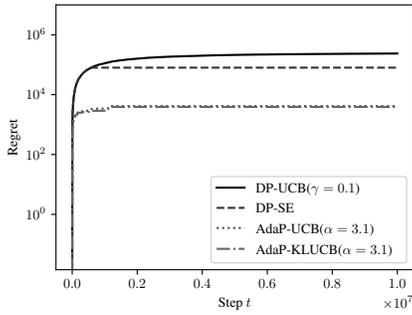
Figure 6: Evolution of regret over time for DP-UCB, DP-SE, AdaP-UCB, and AdaP-KLUCB under C_3 for different values of the privacy budget ϵ . AdaP-KLUCB achieves the lowest regret.



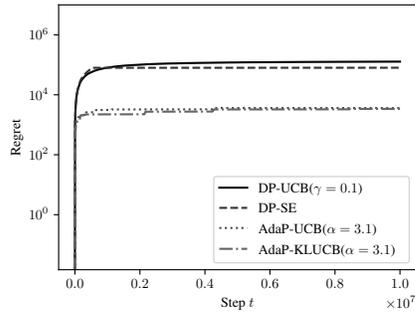
(a) $\epsilon = 0.1$



(b) $\epsilon = 0.25$



(c) $\epsilon = 0.5$



(d) $\epsilon = 1$

Figure 7: Evolution of regret over time for DP-UCB, DP-SE, AdaP-UCB, and AdaP-KLUCB under C_4 for different values of the privacy budget ϵ . AdaP-KLUCB achieves the lowest regret.

G Extended Experimental Analysis

G.1 Experimental Setup

In this section, we perform additional experiments to compare AdaP-UCB and AdaP-KLUCB with respect to the existing bandit algorithms satisfying global DP, i.e. DP-SE (Sajed and Sheffet, 2019) and DP-UCB (Mishra and Thakurta, 2015). We test the four algorithms in the four bandit environments with Bernoulli distributions, as defined by Sajed and Sheffet (2019), namely

$$C_1 = \{0.75, 0.70, 0.70, 0.70, 0.70\}, \quad C_2 = \{0.75, 0.625, 0.5, 0.375, 0.25\}, \\ C_3 = \{0.75, 0.53125, 0.375, 0.28125, 0.25\}, \quad C_4 = \{0.75, 0.71875, 0.625, 0.46875, 0.25\}.$$

For each bandit environment, we implement the algorithms with $\epsilon \in \{0.1, 0.25, 0.5, 1\}$. We set $\alpha = 3.1$ to comply with the regret upper bounds of AdaP-UCB and AdaP-KLUCB. We set $\gamma = 0.1$ for DP-UCB and $\beta = 1/T$. All the algorithms are implemented in Python (version 3.8) and are tested with an 8 core 64-bits Intel i5@1.6 GHz CPU. We run each algorithm 20 times, and plot their average regrets over the runs in Figures 4, 5, 6, and 7. In Section 5, we include Figure 2 to illustrate the evolution of the regret for the four algorithms with environment C_2 and $\epsilon = 1$.

G.2 Experimental Results

Here, we summarise the observations obtained from the experimental results.

Comparative Performance. All the experiments validate that AdaP-KLUCB is the most optimal algorithm satisfying ϵ -global DP for stochastic bandits. Both AdaP-UCB and AdaP-KLUCB achieve similar regret, but AdaP-KLUCB is slightly better in all the cases studied. This observation matches the proven upper bounds, and also reflects similar improvement that KL-UCB brings over UCB in non-private bandits.

Dependence of Regret on ϵ . As predicted by the theoretical analysis, AdaP-UCB and AdaP-KLUCB have different regret depending on ϵ : the regret is smaller for low-privacy regimes. This is also the case for DP-UCB. However, DP-SE have the same performance for different choices of ϵ and echoes the experimental results presented in the original paper (Sajed and Sheffet, 2019).

The Shapes of the Regrets. DP-UCB has a regret shaped like the regret of the classic UCB algorithm. The algorithm chooses a different action at each time-step allowing it to still choose exploratory actions. On the other hand, due to the successive elimination, DP-SE "commits" at a certain step to one action (the optimal action with high probability). Thus, the shape of regret for DP-SE is piece-wise linear. On the other hand, AdaP-UCB and AdaP-KLUCB are a trade-off of both strategies: *due to the doubling, both algorithms "commit" for long episodes to near-optimal actions, while still explore the sub-optimal actions for short episodes.*

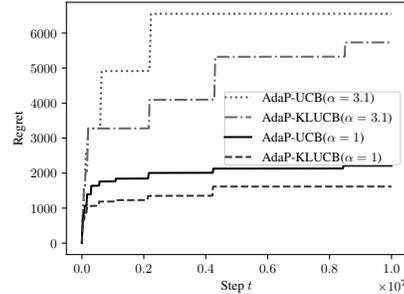


Figure 8: Evolution of regret over time for AdaP-UCB and AdaP-KLUCB for different values of α with C_1 and $\epsilon = 1$. $\alpha = 1$ performs better.

G.3 Choice of α

α controls the width of the optimistic confidence bound. Specifically, it dictates that the real mean is smaller than the optimistic index with high probability, i.e. with probability $1 - \frac{1}{t^\alpha}$ at step t . The requirement that $\alpha > 3$ is due to our analysis of the algorithm. To be specific, the requirement that $\alpha > 3$ is needed to use a sum-integral inequality to bound Term 2 of Step 3 in the proof of Theorem 11. We leave it for future work to relax this requirement.

The experiments are done with $\alpha = 3.1$ to comply with the theoretical analysis. As shown in Figure 8, choosing $\alpha = 1$ works better experimentally. This observation complies with the theoretical results, since the dominant terms in the regret upper bounds of both AdaP-UCB and AdaP-KLUCB are multiplicative in α . A tighter analysis might give us a bound for $\alpha = 1$ and close the multiplicative gap between the regret's lower and upper bound. Reflecting this phenomenon in the analysis will be an interesting future work to pursue.