

810 811 A RELATED WORK

812 813 814 In the context of distributed learning with label shifts, importance ratio estimation is tackled either by solving a linear system as in [\(Lipton et al., 2018;](#page-0-3) [Azizzadenesheli et al., 2019\)](#page-0-4) or by minimizing distribution divergence as in $\sqrt{Garg \text{ et al.} \vert 2020}$. In this section, we overview complete related work.

816 817 818 819 820 821 822 823 824 825 826 827 828 829 830 831 832 833 834 Federated learning (FL). Much of the current research in FL predominantly centers around the minimization of empirical risk, operating under the assumption that each node maintains the same training/test data distribution [\(Li et al., 2020a;](#page-0-6) [Kairouz et al., 2021;](#page-0-7) [Wang et al., 2021b\)](#page-0-8). Prominent methods in FL [\(Kairouz et al., 2021;](#page-0-7) [Li et al., 2020a;](#page-0-6) [Wang et al., 2021b\)](#page-0-8) include FedAvg [\(McMahan](#page-0-9) et al.<mark>,</mark> 2017), FedBN [\(Li et al., 2021b\)](#page-0-10), FedProx [\(Li et al., 2020b\)](#page-0-11) and SCAFFOLD [\(Karimireddy et al.,](#page-0-12) $2020a$). FedAvg and its variants such as $(Huang et al.]$ 2021 ; Karimireddy et al., $2020b$) have been the subject of thorough investigation in optimization literature, exploring facets such as communication efficiency, node participation, and privacy assurance [\(Ramezani-Kebrya et al., 2023\)](#page-0-15).Subsequent work, such as the study by $\left|\frac{dE}{dt}\right| \leq \frac{dE}{dt}$ [\(2022\)](#page-0-16), explores Federated Domain Generalization and introduces data augmentation to the training. This model aims to generalize to both in-domain datasets from participating nodes and an out-of-domain dataset from a non-participating node. Additionally, [Gupta et al.](#page-0-17) [\(2022\)](#page-0-17) introduces FL Games, a game-theoretic framework designed to learn causal features that remain invariant across nodes. This is achieved by employing ensembles over nodes' historical actions and enhancing local computation, under the assumption of consistent training/test data distribution across nodes. The existing strategies to address statistical heterogeneity across nodes during training primarily rely on heuristic-based personalization methods, which currently lack theoretical backing in statistical learning $(Smith et al., 2017; Khodak et al., 2019; Li et al., 2021a).$ $(Smith et al., 2017; Khodak et al., 2019; Li et al., 2021a).$ $(Smith et al., 2017; Khodak et al., 2019; Li et al., 2021a).$ $(Smith et al., 2017; Khodak et al., 2019; Li et al., 2021a).$ In contrast, we aim to minimize overall test error amid both intra-node and inter-node distribution shifts, a situation frequently observed in real-world scenarios. Techniques ensuring communication efficiency, robustness, and secure aggregations serve as complementary.

835 836 837 838 839 840 841 Importance ratio estimation Classical Empirical Risk Minimization (ERM) seeks to minimize the expected loss over the training distribution using finite samples. When faced with distribution shifts, the goal shifts to minimizing the expected loss over the target distribution, leading to the development of Importance-Weighted Empirical Risk Minimization (IW-ERM)[\(Shimodaira, 2000;](#page-0-21) [Sugiyama et al., 2006;](#page-0-22) [Byrd & C. Lipton, 2019;](#page-0-23) [Fang et al., 2020\)](#page-0-24). [Shimodaira](#page-0-21) [\(2000\)](#page-0-21) established that the IW-ERM estimator is asymptotically unbiased. Moreover, [Ramezani-Kebrya et al.](#page-0-15) [\(2023\)](#page-0-15) introduced FTW-ERM, which integrates density ratio estimation.

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843 844 845 846 847 848 849 Label shift and MLLS family For theoretical analysis, the conditional distribution $p(x|y)$ is held strictly constant across all distributions [\(Lipton et al., 2018;](#page-0-3) [Garg et al., 2020;](#page-0-5) [Saerens et al., 2002\)](#page-0-25). Both BBSE [\(Lipton et al., 2018\)](#page-0-3) and RLLS [\(Azizzadenesheli et al., 2019\)](#page-0-4) designate a discrete latent space *z* and introduce a confusion matrix-based estimation method to compute the ratio *w* by solving a linear system (Saerens et al., 2002 ; Lipton et al., 2018). This approach is straightforward and has been proven consistent, even when the predictor is not calibrated. However, its subpar performance is attributed to the information loss inherent in the confusion matrix [\(Garg et al., 2020\)](#page-0-5).

850 851 852 853 Consequently, MLLS [\(Garg et al., 2020\)](#page-0-5) introduces a continuous latent space, resulting in a significant enhancement in estimation performance, especially when combined with a post-hoc calibration method [\(Shrikumar et al., 2019\)](#page-0-26). It also provides a consistency guarantee with a canonically calibrated predictor. This EM-based MLLS method is both concave and can be solved efficiently.

854 855 856 857 Discrepancy Measure In information theory and statistics, discrepancy measures play a critical role in quantifying the differences between probability distributions. One such measure is the Bregman Divergence [\(Banerjee et al., 2005\)](#page-0-27), defined as

$$
D_{\phi}(\boldsymbol{x} \| \boldsymbol{y}) = \phi(\boldsymbol{x}) - \phi(\boldsymbol{y}) - \langle \nabla \phi(\boldsymbol{y}), \boldsymbol{x} - \boldsymbol{y} \rangle,
$$

858 859 860 which encapsulates the difference between the value of a convex function ϕ at two points and the value of the linear approximation of ϕ at one point, leveraging the gradient at another point.

861 862 863 Discrepancy measures are generally categorized into two main families: Integral Probability Metrics (IPMs) and *f*-divergences. IPMs, including Maximum Mean Discrepancy [\(Gretton et al., 2012\)](#page-0-28) and Wasserstein distance [\(Villani, 2009\)](#page-0-29), focus on distribution differences $P - Q$. In contrast, f -divergences, such as KL-divergence [\(Kullback & Leibler, 1951\)](#page-0-30) and Total Variation distance, operate on ratios *P/Q* and do not satisfy the triangular inequality. Interconnections and variations between these families are explored in studies like (f, Γ) -Divergences [\(Birrell et al., 2022\)](#page-0-31), which interpolate between f -divergences and IPMs, and research outlining optimal bounds between them (Agrawal $\&$ [Horel, 2020\)](#page-0-32).

 MLLS [\(Garg et al., 2020\)](#page-0-5) employs f-divergence, notably the KL divergence, which is not a metric as it doesn't satisfy the triangular inequality, and requires distribution *P* to be absolutely continuous with respect to *Q*. Concerning IPMs, while MMD is reliant on a kernel function, it can suffer from the curse of dimensionality when faced with high-dimensional data. On the other hand, the Wasserstein distance can be reformulated using Kantorovich-Rubinstein duality [\(Dedecker et al., 2006;](#page-0-33) [Arjovsky](#page-0-34) l[et al., 2017\)](#page-0-34) as a maximization problem subject to a Lipschitz constrained function $\overline{f} : \mathbb{R}^d \to \mathbb{R}$.

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B BBSE AND MLLS FAMILY

In this section, we summarize the contributions of BBSE [\(Lipton et al., 2018\)](#page-0-3) and MLLS [\(Garg et al.,](#page-0-5) [2020\)](#page-0-5). Our objective is to estimate the ratio $p^{te}(y)/p^{tr}(y)$. We consider a scenario with \overline{m} possible label classes, where $y = c$ for $c \in [m]$. Let $r^* = [r_1^*, \ldots, r_m^*]^\top$ represent the true ratios, with each r_c^* defined as $r_c^* = \frac{p^{\text{te}}(y=c)}{p^{\text{tr}}(y=c)}$ [\(Garg et al., 2020\)](#page-0-5). We then define a family of distributions over \mathcal{Z} , parameterized by $r = [r_1, \ldots, r_m]^\top \in \mathbb{R}^m$, where r_c is the *c*-th element of the ratio vector.

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 $p_{\bm r}(\bm z) := \sum^m_{ }$ $\sum_{c=1} p^{\text{te}}(z|y=c) \cdot p^{\text{tr}}(y=c) \cdot r_c$ (9)

Here, $r_c \ge 0$ for $c \in [m]$ and $\sum_{c=1}^m r_c \cdot p^{\text{tr}}(y = c) = \sum_{c=1}^m p^{\text{te}}(y = c) = 1$ as constraints. When $r = r^*$, e.g., $r_c = r_c^*$ for $c \in [m]$, we have $p_r(z) = p_{r^*}(z) = p^{te}(z)$ [\(Garg et al., 2020\)](#page-0-5). So our task is to find *r* such that

$$
\sum_{c=1}^{m} p^{\text{te}}(z|y=c) \cdot p^{\text{tr}}(y=c) \cdot r_c \boldsymbol{x}
$$

=
$$
\sum_{c=1}^{m} p^{\text{tr}}(z, y=c) \cdot r_c = p^{\text{te}}(z)
$$
 (10)

938 939 940 941 942 943 [Lipton et al.](#page-0-3) [\(2018\)](#page-0-3) introduced Black Box Shift Estimation (BBSE) to address this issue. With a pre-trained classifier f for the classification task, BBSE assumes that the latent space $\mathcal Z$ is discrete and defines $p(z|x) = \delta_{\arg \max f(x)}$, where the output of $f(x)$ is a probability vector (or a simplex) over *m* classes. BBSE estimates $p^{\text{te}}(z|y)$ as a confusion matrix, using both the training and validation data. It calculates $p^{\text{tr}}(y = c)$ from the training set and $p^{\text{te}}(z)$ from the test data. The problem then reduces to solving the following equation:

$$
Aw = B \tag{11}
$$

947 948 949 where $|\mathcal{Z}| = m$, $\mathbf{A} \in \mathbb{R}^{m \times m}$ with $A_{jc} = p^{te}(z = j | y = c) \cdot p^{tr}(y = c)$, and $\mathbf{B} \in \mathbb{R}^m$ with $B_j = p^{\text{te}}(z = j)$ for $c, j \in [m]$.

950 The estimation of the confusion matrix in terms of $p^{te}(z|y)$ leads to the loss of calibration information Garg et al., $[2020]$. Furthermore, when defining $\mathcal Z$ as a continuous latent space, the confusion matrix becomes intractable since *z* has infinitely many values. Therefore, MLLS directly minimizes the divergence between $p^{\text{te}}(z)$ and $p_r(z)$, instead of solving the linear system in Equation [\(11\)](#page-3-1).

Within the *f*-divergence family, MLLS seeks to find a weight vector *r* by minimizing the KLdivergence $D_{\text{KL}}(p^{\text{te}}(z), p_r(z)) = \mathbb{E}_{\text{te}}[\log p^{\text{te}}(z)/p_r(z)]$, for $p_r(z)$ defined in Equation [\(9\)](#page-3-2). Leveraging on the properties of the logarithm, this is equivalent to maximizing the log-likelihood: $r := \arg \max_{r \in \mathbb{R}} \mathbb{E}_{\text{te}} [\log p_r(z)]$. Expanding $p_r(z)$, we have

$$
\mathbb{E}_{\text{te}}\left[\log p_r(z)\right] = \mathbb{E}_{\text{te}}\left[\log\left(\sum_{c=1}^m p^{\text{tr}}(z, y = c)r_c\right)\right]
$$

$$
= \mathbb{E}_{\text{te}}\left[\log\left(\sum_{c=1}^m p^{\text{tr}}(y = c \mid z)r_c\right) + \log p^{\text{tr}}(z)\right].
$$
(12)

Therefore the unified form of MLLS can be formulated as:

$$
\boldsymbol{r} := \arg \max_{\boldsymbol{r} \in \mathbb{R}} \mathbb{E}_{\text{te}} \left[\log \left(\sum_{c=1}^{m} p^{\text{tr}} (y = c \mid \boldsymbol{z}) r_c \right) \right]. \tag{13}
$$

969 970 971 This is a convex optimization problem and can be solved efficiently using methods such as EM, an analytic approach, and also iterative optimization methods like gradient descent with labeled training data and unlabeled test data. MLLS defines the $p(z|x)$ as δ_x , plugs in the pre-defined f to approximate $p^{\text{tr}}(y|\mathbf{x})$ and optimizes the following objective:

 $\boldsymbol{r}_f := \argmax_{\boldsymbol{r} \in \mathbb{R}} \ell(\boldsymbol{r}, f) := \argmax_{\boldsymbol{r} \in \mathbb{R}}$ $r \in \mathbb{R}$ $\mathbb{E}_{\text{te}}\left[\log(f(\boldsymbol{x})^T\boldsymbol{r})\right]$ *.* (14)

With the Bias-Corrected Calibration (BCT) [\(Shrikumar et al., 2019\)](#page-0-26) strategy, they adjust the logits $\hat{f}(\mathbf{x})$ of $f(\mathbf{x})$ element-wise for each class, and the objective becomes:

$$
\boldsymbol{r}_f := \underset{\boldsymbol{r} \in \mathbb{R}}{\arg \max} \ \ell(\boldsymbol{r}, f) := \underset{\boldsymbol{r} \in \mathbb{R}}{\arg \max} \ \mathbb{E}_{\text{te}} \left[\log(g \circ \hat{f}(\boldsymbol{x}))^T \boldsymbol{r} \right], \tag{15}
$$

where *g* is a calibration function.

1026 1027 1028 1029 1030 1031 1032 Scenario **1988** + 1995 + 1996 + 1996 + 1997 + 1998 + 1 $No-LS$ in equation 16 2_q $p_1^{\text{tr}}(\boldsymbol{y}) = p_1^{\text{te}}(\boldsymbol{y})$ and $p_1^{\text{tr}}(\boldsymbol{y}) \neq p_2^{\text{tr}}(\boldsymbol{y})$ *p*₁ $\frac{\text{tr}}{1}(\boldsymbol{y})/p_2^{\text{tr}}(\boldsymbol{y})$ LS on single in equation [17](#page-5-2) 2 $p_1^{\text{te}}(\bm{y}) \neq p_1^{\text{te}}(\bm{y})$ and $p_2^{\text{te}}(\bm{y}) = p_2^{\text{te}}(\bm{y}) \,\, p_1^{\text{te}}(\bm{y})/p_1^{\text{tr}}(\bm{y})$ and $p_1^{\text{te}}(\bm{y})/p_2^{\text{tr}}(\bm{y})$ LS on both in equation 17 2 $p_1^{\text{te}}(\boldsymbol{y}) \neq p_1^{\text{te}}(\boldsymbol{y})$ and $p_2^{\text{te}}(\boldsymbol{y}) \neq p_2^{\text{te}}(\boldsymbol{y})$ $p_1^{\text{te}}(\boldsymbol{y})/p_1^{\text{tr}}(\boldsymbol{y})$ and $p_1^{\text{te}}(\boldsymbol{y})/p_2^{\text{tr}}(\boldsymbol{y})$ LS on multi in equation 18 *K* $p_1^{\text{tr}}(\mathbf{y}) \neq p_1^{\text{te}}(\mathbf{y})$ for all *k p*^{te} $\int_1^{\text{te}}(\bm{y})/p^{\text{tr}}_k(\bm{y})$ for all k

Table 4: Details of scenarios described in Section [2](#page-0-35)

1035 1036 C PROOF OF PROPOSITION [2.1](#page-0-36)

In the following, we consider four typical scenarios under various distribution shifts and formulate their IW-ERM with a focus on minimizing *R*1.

1042 C.1 NO INTRA-NODE LABEL SHIFT

1043 1044 1045 1046 1047 For simplicity, we assume that there are only 2 nodes, but our results can be extended to multiple nodes. This scenario assumes $p_k^{\text{tr}}(\mathbf{y}) = p_k^{\text{te}}(\mathbf{y})$ for $k = 1, 2$, but $p_1^{\text{tr}}(\mathbf{y}) \neq p_2^{\text{tr}}(\mathbf{y})$. Node 1 aims to learn h_w assuming $\frac{p_1^{\text{tr}}(\bm{y})}{p_2^{\text{tr}}(\bm{y})}$ is given. We consider the following IW-ERM that is consistent in minimizing R_1 :

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$$
\begin{array}{c} 1049 \\ 1050 \end{array}
$$

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1052 1053 $\min_{h_w \in \mathcal{H}}$ 1 $n_1^{\rm tr}$ $\sum^{n_1^{\text{tr}}}$ *i*=1 $\ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{1,i}^{\text{tr}}), \boldsymbol{y}_{1,i}^{\text{tr}})$ $+\frac{1}{1}$ $n_2^{\rm tr}$ $\sum^{n_2^{\text{tr}}}$ *i*=1 $p_1^{\textrm{tr}}(\bm{y}_{2,i}^{\textrm{tr}})$ $p_2^{\text{tr}}(\bm{y}_{2,i}^{\text{tr}})$ $\ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{2,i}^{\text{tr}}), \boldsymbol{y}_{2,i}^{\text{tr}}).$ (16)

 $\ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{1,i}^{\text{tr}}), \boldsymbol{y}_{1,i}^{\text{tr}})$

 $\ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{2,i}^{\text{tr}}), \boldsymbol{y}_{2,i}^{\text{tr}}).$

(17)

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Here H is the hypothesis class of h_w . This scenario is referred to as $No-LS$.

 $\min_{h_w \in \mathcal{H}}$

1 $n_1^{\rm tr}$

 $+\frac{1}{4}$ $n_2^{\rm tr}$

 $\sum^{n_1^{\text{tr}}}$

i=1

 $\sum^{n_2^{\text{tr}}}$

i=1

1059 C.2 LABEL SHIFT ONLY FOR NODE 1

1061 1062 Here we consider label shift only for node 1, i.e., $p_1^{\text{tr}}(\mathbf{y}) \neq p_1^{\text{te}}(\mathbf{y})$ and $p_2^{\text{tr}}(\mathbf{y}) = p_2^{\text{te}}(\mathbf{y})$. We consider the following IW-ERM:

> $p_1^{\text{te}}(\bm{y}_{1,i}^{\text{tr}})$ $p_1^{\textrm{tr}}(\bm{y}_{1,i}^{\textrm{tr}})$

> > $p_1^{\text{te}}(\bm{y}_{2,i}^{\text{tr}})$ $p_2^{\text{tr}}(\bm{y}_{2,i}^{\text{tr}})$

$$
\frac{1063}{1064}
$$

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1072

$$
1070\\
$$

1071 This scenario is referred to as LS on single.

1073 1074 C.3 LABEL SHIFT FOR BOTH NODES

1075 1076 1077 Here we assume $p_1^{\text{tr}}(\mathbf{y}) \neq p_1^{\text{te}}(\mathbf{y})$ and $p_2^{\text{tr}}(\mathbf{y}) \neq p_2^{\text{te}}(\mathbf{y})$, i.e., label shift for both nodes. The corresponding IW-ERM is the same as Eq. equation $\boxed{17}$. This scenario is referred to as LS on both.

1078 1079 Without loss of generality and for simplicity, we set $l = 1$. We consider four typical scenarios under various distribution shifts and formulate their IW-ERM with a focus on minimizing *R*1. The details of these scenarios are summarized in Table $\overline{4}$.

1080 1081 C.4 MULTIPLE NODES

1082 1083 Here we consider a general scenario with *K* nodes. We assume both intra-node and inter-node label shifts by the following IW-ERM:

$$
\min_{h_{\mathbf{w}} \in \mathcal{H}} \sum_{k=1}^{K} \frac{\lambda_k}{n_k^{\text{tr}}} \sum_{i=1}^{n_k^{\text{tr}}} \frac{p_1^{\text{te}}(\mathbf{y}_{k,i}^{\text{tr}})}{p_k^{\text{tr}}(\mathbf{y}_{k,i}^{\text{tr}})} \ell(h_{\mathbf{w}}(\mathbf{x}_{k,i}^{\text{tr}}), \mathbf{y}_{k,i}^{\text{tr}}), \tag{18}
$$

(19)

1087 1088 where $\sum_{k=1}^{K} \lambda_k = 1$ and $\lambda_k \geq 0$. This scenario is referred to as LS on multi.

For the scenario without intra-node label shift, the IW-ERM in Equation [\(16\)](#page-5-1) can be expressed as

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\n
$$
\frac{1}{n_2^{\text{tr}}} \sum_{i=1}^{n_2^{\text{tr}}} \frac{p_1^{\text{tr}}(\mathbf{y}_{2,i}^{\text{tr}})}{p_2^{\text{tr}}(\mathbf{y}_{2,i}^{\text{tr}})} \ell(h_{\mathbf{w}}(\mathbf{x}_{2,i}^{\text{tr}}), \mathbf{y}_{2,i}^{\text{tr}})
$$
\n
$$
\xrightarrow{n_2^{\text{tr}} \rightarrow \infty} \mathbb{E}_{p_2^{\text{tr}}(\mathbf{x}, \mathbf{y})} \left[\frac{p_1^{\text{tr}}(\mathbf{y})}{p_2^{\text{tr}}(\mathbf{y})} \ell(h_{\mathbf{w}}(\mathbf{x}), \mathbf{y}) \right]
$$
\n
$$
= \int p_1^{\text{tr}}(\mathbf{y})_{\mathbb{F}} \ldots \frac{\ell(h_{\mathbf{w}}(\mathbf{x}, \mathbf{w})_{\mathbf{w}}^{\text{tr}}(\mathbf{y}) d\mathbf{w}}{\ell(h_{\mathbf{w}}(\mathbf{x}, \mathbf{w})_{\mathbf{w}}^{\text{tr}}(\mathbf{y}) d\mathbf{w}}.
$$

1099 1100 1101 1102 = *Y p*tr ^E*p*(*x|y*)[`(*hw*(*x*)*, ^y*)]*p*tr ²(*y*)*dy*) ²(*y*) = Z *Y p*tr ¹(*y*)E*p*(*x|y*)[`(*hw*(*x*)*, y*)]*dy* = *p*te ¹ (*y*)E*p*(*x|y*)[`(*hw*(*x*)*, y*)]*dy*

 $R_1(h_w)$.

1103
\n
$$
\int y^{1+\epsilon/2} e^{i\omega/2} f(\epsilon/2)
$$
\n
$$
= \mathbb{E}_{p_1^{te}}(x, y) \left[\ell(h_{\boldsymbol{w}}(\boldsymbol{x}), y) \right]
$$

*n*tr

$$
1105 = 1106
$$

1107 1108 1109 where the second equality holds due to the assumption of the label shift setting and Bayes' theorem: $p(x, y) = p(x|y) \cdot p(y)$, and the fourth equality holds by the assumption that $p_1^{\text{tr}}(y) = p_1^{\text{te}}(y)$ in the No-LS setting.

1110 1111 1112 For the scenario with label shift only for Node 1 or for both nodes, the IW-ERM in Equation ($\overline{17}$) admits

$$
\frac{1}{n_2^{\text{tr}}} \sum_{i=1}^{n_2^{\text{tr}}} \frac{p_1^{\text{te}}(\boldsymbol{y}_{2,i}^{\text{tr}})}{p_2^{\text{tr}}(\boldsymbol{y}_{2,i}^{\text{tr}})} \ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{2,i}^{\text{tr}}), \boldsymbol{y}_{2,i}^{\text{tr}})
$$
(20)

$$
\xrightarrow{n_{2}^{\text{tr}} \to \infty} \mathbb{E}_{p_{2}^{\text{tr}}(\boldsymbol{x}, \boldsymbol{y})} \left[\frac{p_{1}^{\text{te}}(\boldsymbol{y})}{p_{2}^{\text{tr}}(\boldsymbol{y})} \ell(h_{\boldsymbol{w}}(\boldsymbol{x}), \boldsymbol{y}) \right]
$$
(21)

$$
= \int_{\mathcal{Y}} \frac{p_1^{\text{te}}(y)}{p_2^{\text{te}}(y)} \mathbb{E}_{p(\boldsymbol{x}|\boldsymbol{y})} [\ell(h_{\boldsymbol{w}}(\boldsymbol{x}), \boldsymbol{y})] p_2^{\text{te}}(\boldsymbol{y}) d\boldsymbol{y} \tag{22}
$$

$$
1120
$$

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$$
= \int_{\mathcal{Y}} p_1^{\text{te}}(y = y) \mathbb{E}_{p(\boldsymbol{x}|\boldsymbol{y})} [\ell(h_{\boldsymbol{w}}(\boldsymbol{x}), y)] d\boldsymbol{y}
$$
(23)

$$
= \mathbb{E}_{p_1^{\text{te}}(\boldsymbol{x}, \boldsymbol{y})} \left[\ell(h_{\boldsymbol{w}}(\boldsymbol{x}), \boldsymbol{y}) \right] \tag{24}
$$

$$
=R_1(h_{\mathbf{w}}). \tag{25}
$$

1126 For multiple nodes, let $k \in [K]$. Similarly, we have

$$
\frac{1}{n_k^{\text{tr}}} \sum_{i=1}^{n_k^{\text{tr}}} \frac{p_1^{\text{te}}(\mathbf{y}_{k,i}^{\text{tr}})}{p_k^{\text{tr}}(\mathbf{y}_{k,i}^{\text{tr}})} \ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{k,i}^{\text{tr}}), \mathbf{y}_{k,i}^{\text{tr}}) \xrightarrow{n_k^{\text{tr}} \to \infty} R_1(h_{\boldsymbol{w}}). \tag{26}
$$

1130 Then we have

1131 Then we have
\n
$$
\sum_{k=1}^{K} \frac{\lambda_k}{n_k^{\text{tr}}} \sum_{i=1}^{n_k^{\text{tr}}} \frac{p_1^{\text{te}}(\boldsymbol{y}_{k,i}^{\text{tr}})}{p_k^{\text{te}}(\boldsymbol{y}_{k,i}^{\text{tr}})} \ell(h_{\boldsymbol{w}}(\boldsymbol{x}_{k,i}^{\text{tr}}), \boldsymbol{y}_{k,i}^{\text{tr}}) \xrightarrow{n_1^{\text{tr}}, \dots, n_K^{\text{tr}} \to \infty} R_1(h_{\boldsymbol{w}}).
$$
\n(27)

 D ALGORITHMIC DESCRIPTION

Algorithm 3 IW-ERM with VRLS in Distributed Learning **Require:** Labeled training data $\{(\boldsymbol{x}_{k,i}^{\text{tr}}, \boldsymbol{y}_{k,i}^{\text{tr}})\}_{i=1}^{n_k^{\text{tr}}}$ at each node *k*, for $k = [K]$. **Require:** Unlabeled test data $\{x_{k,j}^{\text{te}}\}_{j=1}^{n_k^{\text{te}}}$ at each node *k*, for $k = [K]$. Require: Initial global model *hw*. Ensure: Trained global model *h^w* optimized with IW-ERM. 1: Phase 1: Density Ratio Estimation with VRLS 2: for each node $k = 1$ to K in parallel do 3: Train local predictor $f_{k, \hat{\theta}_{n_x}^{\text{tr}}}$ on local training data $\{(\boldsymbol{x}_{k,i}^{\text{tr}}, \boldsymbol{y}_{k,i}^{\text{tr}})\}.$ 4: Use $f_{k, \hat{\theta}_{n_k^{\text{tr}}}}$ to estimate the density ratio $\hat{r}_{n_k^{\text{te}}}$ on unlabelled test data $\{\boldsymbol{x}_k^{\text{te}}\}$ at node *k*. 5: end for 6: Phase 2: Importance Weight Computation 7: **for each node** $k = 1$ to K **do**
8: Compute importance weig Compute importance weight: $\omega_k =$ $\sum_{j=1}^K \hat{\bm{r}}_{n_j^\textsf{te}} \cdot p_j^\textsf{tr}(\bm{y})$ $p_k^{\text{tr}}(\bm{y})$ 9: end for 10: Phase 3: Global Model Training with IW-ERM 11: Train global model h_w by minimizing the weighted empirical risk: min *hw* X*K k*=1 λ_k n_k^{tr} \sum_{k} $\sum_{i=1}^N \omega_k \cdot \ell\left(h_{\boldsymbol{w}}(\boldsymbol{x}_{k,i}^{\text{tr}}), \boldsymbol{y}_{k,i}^{\text{tr}}\right)$

```
1242
1243
1
2 # Split the training dataset on each node
1244
3 trainsets = target_shift.split_dataset(trainset.data, trainset.targets,
1245
1246
4
<sup>1247</sup> <sup>5</sup> # Split the test dataset on each node
1248
1249
7
1250
8 # Initialize K local models (nets) for each node
1251
9 nets = [initialize_model() for _ in range(node_num)]
1252
10
1253
12 estimators = [LS_RatioModel(nets[k]) for k in range(node_num)]
1254
13
1255
14 # Initialize tensors to store the estimated ratios, values, and marginal
1256
1257
15 estimated_ratios = torch.zeros(node_num, node_num, nclass)
1258
17 marginal_values = torch.zeros(node_num, nclass)
1259
18
1260
19 # Phase 1: Compute the estimated ratios for each node pair (k, j)
1261
20 for k in range(node_num):
1262<sup>21</sup><sub>22</sub>
1263
1264
1265
24
1266
<sup>1267</sup><sub>26</sub> for i, trainset in enumerate(trainsets):<br>1267<sub>26</sub> marginal upluse[i] = marginal(trains
1268
28
1269
29 # Phase 3: Compute the final estimated values for each node
1270
30 for k in range(node_num):
1271
1272
1273
33
1274
34 # Aggregate the estimated values across nodes
1275
35 aggregated_values = torch.sum(estimated_values, dim=1)
1276
36
1277<sup>37</sup> # Compute the final ratios for each node
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          node label dist train, transform=transform train)
    6 testsets = target_shift.split_dataset(testset.data, testset.targets,
           node_label_dist_test, transform=transform_test)
    11 # Initialize the estimator for each local model
           values for each pair of nodes.
    16 estimated_values = torch.zeros(node_num, node_num, nclass)
          for j in range(node_num):
               # Perform test on node k using node j's testset
               estimated_ratios[k, j] = estimators[k](testsets[j].data.cpu().
          numpy())
1266^{25} # Phase 2: Compute the marginal values on each node's training set
           marginal_values[i] = marginal(trainset.targets)
           for j in range(node_num):
                estimated_values[k, j] = marginal_values[j] \star estimated_ratios[k,
            j]
    38 ratios = (aggregated_values / marginal_values).to(args.device)
             Listing 1: Our VRLS in distributed learning. It is the implementation of Algorithm \beta
```
1296 1297 E PROOF OF THEOREM [5.1](#page-0-0)

1298 1299 *Proof.* Let $H(r, \theta, x) = -\log(f(x, \theta)^T r)$. From the strong convexity in Lemma \mathbb{E} .7, we have that

$$
\|\hat{r}_{n^{te}} - r_{f^*}\|_2^2 \le \frac{2}{\mu p_{\min}} \left(\mathcal{L}_{\theta^*}(\hat{r}_{n^{te}}) - \mathcal{L}_{\theta^*}(r_{f^*}) \right) \tag{28}
$$

 $+2L\mathbb{E}$

Now focusing on the term on the right-hand side, we find by invoking Lemma [E.4](#page-11-0) that

 $-\mathbb{E}$

$$
\mathcal{L}_{{\bm{\theta}}^\star}(\hat{\bm{r}}_{n^\text{te}})-\mathcal{L}_{{\bm{\theta}}^\star}(\bm{r}_{f^\star})
$$

$$
\leq \mathbb{E}\bigg[H(\hat{\boldsymbol{r}}_{n^{\text{te}}},\hat{\boldsymbol{\theta}}_{n^{\text{tr}}},\boldsymbol{x})\bigg]-\mathbb{E}\bigg[H(\boldsymbol{r}_{f^\star},\hat{\boldsymbol{\theta}}_{n^{\text{tr}}},\boldsymbol{x})\bigg]+2L\mathbb{E}\bigg[\|\hat{\boldsymbol{\theta}}_{n^{\text{tr}}}-\boldsymbol{\theta}^\star\|_2\bigg] \\-\bigg[\cos\left(\frac{2\pi}{\lambda}\right)\bigg]-\frac{1}{n^{\text{te}}}\cos\left(\frac{2\pi}{\lambda}\right)\bigg]+2L\mathbb{E}\bigg[\|\hat{\boldsymbol{\theta}}_{n^{\text{tr}}}-\boldsymbol{\theta}^\star\|_2\bigg]
$$

$$
= \mathbb{E}\bigg[H(\hat{\boldsymbol{r}}_{n^\textup{\text{te}}},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},x)\bigg]-\frac{1}{n^\textup{\text{te}}}\sum_{j=1}^{n^\textup{\text{te}}}H(\hat{\boldsymbol{r}}_{n^\textup{\text{te}}},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x}_j)+\frac{1}{n^\textup{\text{te}}}\sum_{j=1}^{n^\textup{\text{te}}}H(\hat{\boldsymbol{r}}_{n},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x}_j)\\-\mathbb{E}\bigg[H(\boldsymbol{r}_{f^\star},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x})\bigg]+2L\mathbb{E}\bigg[\|\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}}-\boldsymbol{\theta}^\star\|_2\bigg]
$$

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$$
f_{\rm{max}}
$$

$$
\leq \mathbb{E}\bigg[H(\hat{\boldsymbol{r}}_{n^\textup{\text{te}}},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x})\bigg]-\frac{1}{n^\textup{\text{te}}}\sum_{j=1}^{n^\textup{\text{te}}}H(\hat{\boldsymbol{r}}_{n^\textup{\text{te}}},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x}_j)+\frac{1}{n^\textup{\text{te}}}\sum_{j=1}^{n^\textup{\text{te}}}H(\boldsymbol{r}_{f^\star},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x}_j)\\-\mathbb{E}\Big[H(\boldsymbol{r}_{f^\star},\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}},\boldsymbol{x})\Big]+2L\mathbb{E}\Big[\|\hat{\boldsymbol{\theta}}_{n^\textup{\text{tr}}}-\boldsymbol{\theta}^\star\|_2\Big].
$$

$$
-\mathbb{E}\left[H(\boldsymbol{r}_{f^*},\hat{\boldsymbol{\theta}}_{n^{\text{tr}}},\boldsymbol{x})\right]+2L\mathbb{E}\left[\|\hat{\boldsymbol{\theta}}_{n^{\text{tr}}}-\boldsymbol{\theta}^*\|_2\right],
$$
\n(29)

1320 1321 where in the last inequality we used the fact that \hat{r}_n is a minimizer of $r \mapsto \frac{1}{n} \sum_{j=1}^n H(r, \hat{\theta}_t, x_j)$. Finally by using Lemma $\overline{E.5}$ and Lemma $\overline{E.6}$ with $\delta/2$ each, we have that with probability $1 - \delta$,

$$
\mathcal{L}_{\theta^{\star}}(\hat{r}_{n^{te}}) - \mathcal{L}_{\theta^{\star}}(r_{f^{\star}}) \leq \frac{4}{\sqrt{n^{te}}} \text{Rad}(\mathcal{F}) + 2L\mathbb{E}\left[\|\hat{\theta}_{n^{tr}} - \theta^{\star}\|_{2}\right] + 4B\sqrt{\frac{\log(4/\delta)}{n^{te}}}\tag{30}
$$

1325 Plugging this back into Equation (28) , we have that

$$
\|\hat{\boldsymbol{r}}_{n^{\text{te}}} - \boldsymbol{r}_{f^{\star}}\|_{2}^{2} \leq \frac{2}{\mu p_{\min}} \left(\frac{4}{\sqrt{n^{\text{te}}}} \text{Rad}(\mathcal{F}) + 4B \sqrt{\frac{\log(4/\delta)}{n^{\text{te}}}} \right) + \frac{4L}{\mu p_{\min}} \mathbb{E}\left[\|\hat{\boldsymbol{\theta}}_{n^{\text{tr}}} - \boldsymbol{\theta}^{\star}\|_{2} \right].
$$
 (31)

1330 1331 Lemma E.1. *For any* $r \in \mathbb{R}^m_+$, $\theta \in \Theta$, $x \in \mathcal{X}$, we have that

$$
\boldsymbol{r}^\top f(\boldsymbol{x}, \boldsymbol{\theta}) \leq \frac{1}{p_{min}}.
$$

1335 *Proof.* Applying Hölder's inequality we have that

$$
\boldsymbol{r}^\top f(\boldsymbol{x}, \boldsymbol{\theta}) \leq \| \boldsymbol{r} \|_\infty \| f(\boldsymbol{x}, \boldsymbol{\theta}) \|_1 = \| \boldsymbol{r} \|_\infty.
$$

1338 1339 1340 Moreover, since $r \in \mathbb{R}^m_+$, we have that $\sum_y r_y p_{tr}(y) = 1$ This implies that $\|r\|_{\infty} \le \frac{1}{p_{\min}}$, which yields the result.

1341 1342 Lemma E.2 (Implication of Assumption Assumption $\overline{5.1}$). *Under Assumption* $\overline{5.1}$ *, there exists* $B > 0$ *such that for any* $\mathbf{r} \in \mathbb{R}^m_+$, $\mathbf{\theta} \in \Theta$, $\mathbf{x} \in \mathcal{X}$,

$$
|\log(\mathbf{r}^{\top}f(\mathbf{x},\boldsymbol{\theta}))| \leq B.
$$

1343 1344

1345 1346 *Proof.* Since $r \in \mathbb{R}_+^m$, it has at least one non-zero coordinate and $f(x, \theta)$ is the output of a softmax layer so all of its coordinates are non-zero. Consequently,

$$
\mathbf{r}^{\top}f(\mathbf{x},\boldsymbol{\theta})>0
$$

1349 So by Assumption $[5.1]$ the function $(r, \theta, x) \mapsto \log(r^{\top} f(x, \theta))$ is defined and continuous over a compact set, so there exists a constant *B* giving us the result. compact set, so there exists a constant *B* giving us the result.

1350 1351 1352 Lemma E.3 (Population Strong Convexity). Let $H(r, \theta, x) = -\log(r^{\top} f(x, \theta))$. Under Assump*tion Assumption [5.2,](#page-0-39) the function*

$$
\mathcal{L}_{\boldsymbol{\theta}^\star}: \boldsymbol{r} \mapsto \mathbb{E}\bigg[H(\boldsymbol{r}, \boldsymbol{\theta}^\star, \boldsymbol{x})\bigg]
$$

1355 *is µp*min*-strongly convex.*

Proof. We first compute the Hessian of *L* to find that

$$
\nabla^2 \mathcal{L}(\boldsymbol{r}) = \mathbb{E}\bigg[\frac{1}{(\boldsymbol{r}^\top f(\boldsymbol{x}, \boldsymbol{\theta}^\star))^2} f(\boldsymbol{x}, \boldsymbol{\theta}^\star) f(\boldsymbol{x}, \boldsymbol{\theta}^\star)^\top\bigg].
$$

1360 1361 Since by Lemma $\boxed{E.1}$, we have that $r^{\top}f(x, \theta^*) \leq p_{\min}^{-1}$, we conclude that

$$
\nabla^2 \mathcal{L}(\bm{r}) \succeq p_{\min} \mathbb{E}\bigg[f(\bm{x}, \bm{\theta}^{\star}) f(\bm{x}, \bm{\theta}^{\star})^{\top}\bigg] \succeq \mu p_{\min} \mathbf{I}_m.
$$

 \Box

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1365 1366 1367 Lemma E.4 (Lipschitz Parametrization). Let $H(r, \theta, x) = -\log(f(x, \theta)^T r)$. There exists $L > 0$ $\mathit{such that for any } \theta_1, \theta_2 \in \Theta, \textit{and } \mathbf{r} \in \mathbb{R}_+^m, \textit{we have that}$

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1369 1370 *Proof.* The gradient of *H* with respect to θ is given by

$$
\nabla_{\boldsymbol{\theta}} H(\boldsymbol{r}, \boldsymbol{\theta}, \boldsymbol{x}) = -\frac{1}{f(\boldsymbol{x}, \boldsymbol{\theta})^\top \boldsymbol{r}} \nabla_{\boldsymbol{\theta}} f(\boldsymbol{x}, \boldsymbol{\theta})
$$

 $|H(r, \theta_1, x) - H(r, \theta_2, x)| \leq L \|\theta_1 - \theta_2\|_2.$

1373 1374 1375 1376 Reasoning like in Lemma $E.1$, we know that $\frac{1}{f(x,\theta)^{\top}r}$ is defined and continuous over the compact set of its parameters, we also know that *f* is a neural network parametrized by θ , hence $\nabla_{\theta} f(x, \theta)$ is bounded when θ and x are bounded. Consequently, under Assumption $\overline{5.1}$, there exists a constant $L > 0$ such that $\|\nabla_{\theta}H(\mathbf{r}, \mathbf{A}|\mathbf{r})\|_{2} < I$

$$
\|\mathbf{v}_{\boldsymbol{\theta}}\mathbf{\mu}(\mathbf{r},\mathbf{v},\mathbf{w})\|_2 \geq L.
$$

1379 1380 Lemma E.5 (Uniform Bound 1). Let $\delta \in (0,1)$, with probability $1 - \delta$, we have that

$$
\mathbb{E}\left[H(\hat{\boldsymbol{r}}_n, \hat{\boldsymbol{\theta}}_t, \boldsymbol{x})\right] - \frac{1}{n} \sum_{j=1}^n H(\hat{\boldsymbol{r}}_n, \hat{\boldsymbol{\theta}}_t, \boldsymbol{x}_j) \n\leq \frac{2}{\sqrt{n}} Rad(\mathcal{F}) + 2B\sqrt{\frac{\log(4/\delta)}{n}}.
$$
\n(32)

1387 1388 1389 *Proof.* Let $\delta \in (0, 1)$. Since \hat{r}_n is learned from the samples x_j , we do not have independence, which would have allowed us to apply a concentration inequality. Hence, we derive a uniform bound as follows. We begin by observing that:

$$
\mathbb{E}\bigg[H(\hat{\boldsymbol{r}}_n,\hat{\boldsymbol{\theta}}_t,\boldsymbol{x})\bigg]-\frac{1}{n}\sum_{j=1}^n H(\hat{\boldsymbol{r}}_n,\hat{\boldsymbol{\theta}}_t,\boldsymbol{x}_j)
$$

$$
\begin{aligned}\n\frac{1393}{1394} &\leq \sup_{\boldsymbol{r},\boldsymbol{\theta}} \left(\mathbb{E} \bigg[H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x}) \bigg] - \frac{1}{n} \sum_{j=1}^{n} H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x}_{j}) \right)\n\end{aligned}
$$

1396 1397 Now since Lemma [E.2](#page-10-3) holds, we can apply McDiarmid's Inequality to get that with probability $1 - \delta$, we have:

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\n
$$
\sup_{\boldsymbol{r},\boldsymbol{\theta}} \left(\mathbb{E} \bigg[H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x}) \bigg] - \frac{1}{n} \sum_{j=1}^n H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x}_j) \right)
$$

1401 *j*=1

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\n1403
$$
\leq \mathbb{E}\bigg[\sup_{\boldsymbol{r},\boldsymbol{\theta}}\left(\mathbb{E}\big[H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x})\big] - \frac{1}{n}\sum_{j=1}^n H(\boldsymbol{r},\boldsymbol{\theta},\boldsymbol{x}_j)\right)\bigg] + 2B\sqrt{\frac{\log(2/\delta)}{n}}
$$

1404 1405 1406 The expectation of the supremum on the right-hand side can be bounded by the Rademacher complexity of $\mathcal{F} := \{ \mathbf{x} \mapsto \mathbf{r}^\top f(\mathbf{x}, \mathbf{\theta}), \ (\mathbf{r}, \mathbf{\theta}) \in \mathbb{R}^m_+ \times \Theta \}$, and we obtain:

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\n
$$
\leq \frac{2}{\sqrt{n}} \text{Rad}(\mathcal{F}) + 2B \sqrt{\frac{\log(2/\delta)}{n}}.
$$
\n(33)

1414 1415 Lemma E.6 (Uniform Bound 2). Let $\delta \in (0,1)$, with probability $1 - \delta$, we have that

$$
\mathbb{E}\left[H(\mathbf{r}_{f^*}, \hat{\boldsymbol{\theta}}_t, \mathbf{x})\right] - \frac{1}{n} \sum_{j=1}^n H(\mathbf{r}_{f^*}, \hat{\boldsymbol{\theta}}_t, \mathbf{x}_j)
$$
\n
$$
\leq \frac{2}{\sqrt{n}} Rad(\mathcal{F}) + 2B\sqrt{\frac{\log(2/\delta)}{n}}.
$$
\n(34)

1422 *Proof.* The proof is identical to that of Lemma [E.5.](#page-11-1)

1423 1424 1425 1426 Lemma E.7 (Strong Convexity of Population Loss). Let $\mathcal{L}(r, \theta)$ be the population loss as defined in *Lemma* $\overline{E.7}$. We establish that $\mathcal{L}(r, \theta)$ is μ_{min} -strongly convex under the assumptions of calibration *(Assumption [5.2\)](#page-0-39).*

1428 *Proof.* We compute the Hessian of the population loss $\mathcal L$ as in Lemma $\overline{E.7}$, obtaining that:

$$
\nabla^2 \mathcal{L}(\boldsymbol{r}) = \mathbb{E}\bigg[\frac{1}{(\boldsymbol{r}^\top f(\boldsymbol{x}, \boldsymbol{\theta}))^2} f(\boldsymbol{x}, \boldsymbol{\theta}) f(\boldsymbol{x}, \boldsymbol{\theta})^\top\bigg].
$$

1432 From Lemma $\boxed{E.1}$, we have that $r^{\top}f(x,\theta) \leq p_{\min}^{-1}$. Therefore, we conclude:

$$
\nabla^2 \mathcal{L}(\boldsymbol{r}) \succeq p_{\min} \mathbb{E}\bigg[f(\boldsymbol{x}, \boldsymbol{\theta})f(\boldsymbol{x}, \boldsymbol{\theta})^{\top}\bigg] \succeq \mu p_{\min} \mathbf{I}_m.
$$

1437 1438 1439 Lemma E.8 (Bound on Empirical Loss). *Under Assumption* [5.1,](#page-0-38) *the empirical loss* $\mathcal{L}_{n^{\text{te}}}(\mathbf{r}, \hat{\theta}_{n^{\text{tr}}})$ *satisfies the following concentration bound:*

$$
\mathbb{P}\left(\sup_{\bm{r}\in\mathbb{R}_+^m}\left|\mathcal{L}_{n^{te}}(\bm{r},\hat{\bm{\theta}}_{n^{tr}})-\mathcal{L}(\bm{r},\hat{\bm{\theta}}_{n^{tr}})\right|>\epsilon\right)\leq 2\exp\left(-cn^{te}\epsilon^2\right).
$$

Proof. This result follows from standard concentration inequalities, such as McDiarmid's inequality, together with the Lipschitz continuity of the loss function $\mathcal L$ with respect to the samples. \Box

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 \Box

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1458 1459 1460 F PROOF OF THEOREM **[5.2](#page-0-1)** AND CONVERGENCE-COMMUNICATION GUARANTEES FOR IW-ERM WITH VRLS

1461 1462 1463 1464 1465 1466 1467 1468 1469 We now establish convergence rates for IW-ERM with VRLS and show our proposed importance weighting achieves *the same rates* with the data-dependent *constant terms* increase linearly with $\max_{y \in \mathcal{Y}} \sup_f r_f(y) = r_{\max}$ under negligible communication overhead over the baseline ERMsolvers without importance weighting. In Appendix \overline{F} , we establish tight convergence rates and communication guarantees for IW-ERM with VRLS in a broad range of importance optimization settings including convex optimization, second-order differentiability, composite optimization with proximal operator, optimization with adaptive step-sizes, and nonconvex optimization, along the lines of e.g., [\(Woodworth et al., 2020;](#page-0-30) [Haddadpour et al., 2021;](#page-0-40) [Glasgow et al., 2022;](#page-0-41) [Liu et al., 2023;](#page-0-42) [Hu](#page-0-43) [& Huang, 2023;](#page-0-43) [Wu et al., 2023;](#page-0-44) [Liu et al., 2023\)](#page-0-42).

1470 1471 1472 1473 1474 1475 1476 1477 1478 1479 1480 1481 1482 1483 By estimating the ratios locally and absorbing into local losses, we note that the properties of the modified local loss w.r.t. the neural network parameters *w*, e.g., convexity and smoothness, do not change. The data-dependent parameters such as Lipschitz and smoothness constants for $\ell \circ h_w$ w.r.t. *w* are scaled linearly by r_{max} . Our method of density ratio estimation trains the pre-defined predictor *exclusively using local training data*, which implies IW-ERM with VRLS achieves the same privacy guarantees as the baseline ERM-solvers without importance weighting. For ratio estimation, the communication between clients involves only the estimated marginal label distribution, instead of data, ensuring negligible communication overhead. Given the size of variables to represent marginal distributions, which is by orders of magnitude smaller than the number of parameters of the underlying neural networks for training and the fact that ratio estimation involves only one round of communication, the overall communication overhead for ratio estimation is masked by the communication costs of model training. The communication costs for IW-ERM with VRLS over the course of optimization are exactly the same as those of the baseline ERM-solvers without importance weighting. All in all, importance weighting does not negatively impact communication guarantees throughout the course of optimization, which proves Theorem **5.2**

1484 1485 1486 1487 In the following, we establish tight convergence rates and communication guarantees for IW-ERM with VRLS in a broad range of importance optimization settings including convex optimization, second-order differentiability, composite optimization with proximal operator, optimization with adaptive step-sizes, and nonconvex optimization.

1488 1489 1490 1491 For convex and second-order Differentiable optimization, we establish a lower bound on the convergence rates for IW-ERM in with VRLS and local updating along the lines of e.g., [\(Glasgow et al.,](#page-0-41) 2022 , Theorem 3.1).

1492 1493 1494 1495 1496 1497 Assumption F.1 (PL with Compression). *1) The* $\ell(h_{w}(x), y)$ *is* β -smoothness and convex w.r.t. w *for any* (x, y) *and satisfies Polyak-Łojasiewicz (PL) condition (there exists* $\alpha_{\ell} > 0$ *such that, for all* $w \in W$, we have $\ell(h_w) \leq \|\nabla_w \ell(h_w)\|_2^2/(2\alpha_\ell)$; 2) The compression scheme *Q* is unbiased with *bounded variance, i.e.,* $\mathbb{E}[\mathcal{Q}(x)] = x$ *and* $\mathbb{E}[\|\mathcal{Q}(x) - x\|_2^2 \leq q \|x\|_2^2]$; 3) The stochastic gradient $g(w) = \nabla_w \ell(h_w)$ *is unbiased, i.e.,* $\mathbb{E}[g(w)] = \nabla_w \ell(h_w)$ *for any* $w \in \mathcal{W}$ *with bounded variance* $\mathbb{E}[\|\boldsymbol{g}(\boldsymbol{w})-\nabla_{\boldsymbol{w}}\ell(h_{\boldsymbol{w}})\|_2^2].$

1498 1499 1500 For nonconvex optimization with PL condition and communication compression, we establish convergence and communication guarantees for IW-ERM with VRLS, compression, and local updating along the lines of e.g., $(Haddadpour et al., 2021)$, Theorem 5.1).

1501 1502 1503 1504 1505 1506 Theorem F.1 (Convergence and Communication Bounds for Nonconvex Optimization with PL). *Let* κ *denote the condition number,* τ *denote the number of local steps, R denote the number of communication rounds, and* $\max_{y \in \mathcal{Y}} \sup_f r_f(y) = r_{\max}$ *. Under Assumption F.1*, *suppose Algorithm* [2](#page-0-45)*with* τ *local updates and communication compression [\(Haddadpour et al., 2021,](#page-0-40) Algorithm 1) is run for* $T = \tau R$ *total stochastic gradients per node with fixed step-sizes* $\eta = 1/(2r_{\text{max}}\beta\gamma\tau(q/K + 1))$ *and* $\gamma \geq K$ *. Then we have* $\mathbb{E}[\ell(h_{\mathbf{w}_T}) - \ell(h_{\mathbf{w}^{\star}})] \leq \epsilon$ by setting

$$
R \lesssim \left(\frac{q}{K} + 1\right) \kappa \log\left(\frac{1}{\epsilon}\right) \quad \text{and} \quad \tau \lesssim \left(\frac{q+1}{K(q/K+1)\epsilon}\right). \tag{35}
$$

1509 1510 1511 Assumption F.2 (Nonconvex Optimization with Adaptive Step-sizes). *1) The* $\ell \circ h_w$ *is* β -smoothness *with bounded gradients; 2) The stochastic gradients* $g(w) = \nabla_w \ell(h_w)$ *is unbiased with bounded variance* $\mathbb{E}[\|\mathbf{g}(\mathbf{w}) - \nabla_{\mathbf{w}}\ell(h_{\mathbf{w}})\|_2^2]$; 3) Adaptive matrices A_t constructed as in [\(Wu et al., 2023,](#page-0-44) *Algorithm 2) are diagonal and the minimum eigenvalues satisfy* $\lambda_{\min}(A_t) \ge \rho > 0$ for some $\rho \in \mathbb{R}_+$.

 For nonconvex optimization with adaptive step-sizes, we establish convergence and communication guarantees for IW-ERM with VRLS and local updating along the lines of e.g., \sqrt{Wu} et al., $\sqrt{2023}$, Theorem 2).

 Theorem F.2 (Convergence and Communication Guarantees for Nonconvex Optimization with Adaptive Step-sizes). Let τ denote the number of local steps, R denote the number of communication *rounds, and* $\max_{y \in \mathcal{Y}} \sup_f r_f(y) = r_{\text{max}}$ *. Under Assumption F.2*, *suppose Algorithm* $2 \mid \text{with } \tau \text{ local}$ *updates is run for* $T = \tau R$ *total stochastic gradients per node with an adaptive step-size similar to* (Wu et al., , Algorithm 2). Then we $\mathbb{E}[\|\nabla_{\bm{w}}\ell(h_{\bm{w}_T})\|_2] \leq \epsilon$ by setting:

$$
T \lesssim \frac{r_{\text{max}}}{K\epsilon^3} \quad \text{and} \quad R \lesssim \frac{r_{\text{max}}}{\epsilon^2}.
$$

 Assumption F.3 (Composite Optimization with Proximal Operator). 1) The $\ell \circ h_w$ is smooth and *strongly convex with condition number* κ ; 2) The stochastic gradients $g(w) = \nabla_w \ell(h_w)$ is unbiased.

 For composite optimization with strongly convex and smooth functions and proximal operator, we establish an upper bound on oracle complexity to achieve ϵ error on the Lyapunov function defined as in [\(Hu & Huang, 2023,](#page-0-43) Section 4) for Gradient Flow-type transformation of IW-ERM with VRLS in the limit of infinitesimal step-size.

 Theorem F.3 (Oracle Complexity of Proximal Operator for Composite Optimization). Let κ denote *the condition number. Under Assumption [F.3,](#page-14-0) suppose Gradient Flow-type transformation of Algorithm [2](#page-0-45) with VRLS and Proximal Operator evolves in the limit of infinitesimal step-size [\(Hu &](#page-0-43) Huang* $\frac{2023}{2023}$ *Algorithm 3). Then it achieves* $\mathcal{O}(r_{\text{max}}\sqrt{\kappa} \log(1/\epsilon))$ *Proximal Operator Complexity.*

 G COMPLEXITY ANALYSIS

 In our algorithm, the ratio estimation is performed once in parallel before the IW-ERM step.

 In the experiments, we used a simple network to estimate the ratios in advance, which required significantly less computational effort compared to training the global model. Although IW-ERM with VRLS introduces additional computational complexity compared to the baseline FedAvg, it results in substantial improvements in overall generalization, particularly under challenging label shift conditions.

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 H MATHEMATICAL NOTATIONS

 In this appendix, we provide a summary of mathematical notations used in this paper in Table $5:$

 I LIMITATIONS

 The distribution shifts observed in real-world data are often not fully captured by the label shift or relaxed distribution shift assumptions. In our experiments, we applied mild test data augmentation to approximate the relaxed label shift and manage ratio estimation errors for both the baselines and our method. However, the label shift assumption remains overly restrictive, and the relaxed label shift lacks robust empirical validation in practical scenarios.

 Additionally, IW-ERM's parameter estimation relies on local predictors at each client, which limits its scalability. In practice, a simpler global predictor could be sufficient for parameter estimation and IW-ERM training. Future research could explore VRLS variants capable of effectively handling more complex distribution shifts in challenging datasets, such as CIFAR-10.1 [\(Recht et al., 2018;](#page-0-46) [Torralba](#page-0-3) et al., $[2008]$, as suggested in $(Garg et al.]$ $[2023]$.

1728 1729 J EXPERIMENTAL DETAILS AND ADDITIONAL EXPERIMENTS

1730 1731 1732 In this section, we provide experimental details and additional experiments. In particular, we validate our theory on multiple clients in a federated setting and show that our IW-ERM outperforms FedAvg and FedBN baselines *under drastic and challenging label shifts*.

1734 1735 J.1 EXPERIMENTAL DETAILS

1736 1737 In single-client experiments, a simple MLP without dropout is used as the predictor for MNIST, and ResNet-18 for CIFAR-10.

1738 1739 1740 1741 1742 1743 1744 1745 1746 1747 1748 1749 1750 For experiments in a federated learning setting, both MNIST [\(LeCun et al., 1998\)](#page-0-48) and Fashion MNIST [\(Xiao et al., 2017\)](#page-0-49) datasets are employed, each containing 60,000 training samples and 10,000 test samples, with each sample being a 28 by 28 pixel grayscale image. The CIFAR-10 dataset [\(Krizhevsky\)](#page-0-50) comprises 60,000 colored images, sized 32 by 32 pixels, spread across 10 classes with 6,000 images per class; it is divided into 50,000 training images and 10,000 test images. In this setting, the objective is to minimize the cross-entropy loss. Stochastic gradients for each client are calculated with a batch size of 64 and aggregated on the server using the Adam optimizer. LeNet is used for experiments on MNIST and Fashion MNIST with a learning rate of 0.001 and a weight decay of 1×10^{-6} . For CIFAR-10, ResNet-18 is employed with a learning rate of 0.0001 and a weight decay of 0.0001. Three independent runs are implemented for 5-client experiments on Fashion MNIST and CIFAR-10, while for 10 clients, one run is conducted on CIFAR-10. The regularization coefficient ζ in Equation $\overline{2}$ is set to 1 for all experiments. All experiments are performed using a single GPU on an internal cluster and Colab.

1751 1752 1753 Importantly, the training of the predictor for ratio estimation on both the baseline MLLS and our VRLS is executed with identical hyperparameters and epochs for CIFAR-10 and Fashion MNIST. The training is halted once the classification loss reaches a predefined threshold on MNIST.

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1755 J.2 RELAXED LABEL SHIFT EXPERIMENTS

1756 1757 1758 1759 1760 In conventional label shift, it is assumed that $p(x | y)$ remains unchanged across training and test data. However, this assumption is often too strong for real-world applications, such as in healthcare, where different hospitals may use varying equipment, leading to shifts in $p(x | y)$ even with the same labels [\(Rajendran et al., 2023\)](#page-0-28). Relaxed label shift loosens this assumption by allowing small changes in the conditional distribution [\(Garg et al., 2023;](#page-0-47) [Luo & Ren, 2022\)](#page-0-52).

1761 1762 1763 1764 To formalize this, we use the distributional distance D and a relaxation parameter $\epsilon > 0$, as defined by [Garg et al.](#page-0-47) [\(2023\)](#page-0-47): $\max_{y} \mathcal{D}(p_{tr}(x | y), p_{te}(x | y)) \leq \epsilon$. This allows for slight differences in feature distributions between training and testing, capturing a more realistic scenario where the conditional distribution is not strictly invariant.

1765 1766 1767 1768 1769 In our case, visual inspection suggests that the differences between temporally distinct datasets, such as CIFAR-10 and CIFAR-10.1 v6 [\(Torralba et al., 2008;](#page-0-3) [Recht et al., 2018\)](#page-0-46), may not meet the assumption of a small ϵ . To address this, we instead simulate controlled shifts using test data augmentation, allowing us to regulate the degree of relaxation, following the approach outlined in [Garg et al.](#page-0-47) [\(2023\)](#page-0-47).

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1772 J.3 ADDITIONAL EXPERIMENTS

1773 1774 In this section, we provide supplementary results, visualizations of accuracy across clients and tables showing dataset distribution in FL setting and relaxed label shift.

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VRLS alongside MLLS [\(Garg et al., 2020\)](#page-0-5), EM [\(Saerens et al., 2002\)](#page-0-25), and also RLLS [\(Azizzade](#page-0-4)nesheli et al.,).

Table 6: LeNet on Fashion MNIST with label shift across 5 clients. 15,000 iterations for FedAvg and FedBN; 5,000 for Upper Bound (FTW-ERM) using true ratios and our IW-ERM. To mention, to train our predictor, we use a simpliest MLP and employ linear kernel.

FMNIST	Our IW-ERM	FedAvg	FedBN	Upper Bound		
Avg. accuracy	0.7520 ± 0.0209	$0.5472 + 0.0297$	0.5359 ± 0.0306	0.8273 ± 0.0041		
Client 1 accuracy	0.7162 ± 0.0059	0.3616 ± 0.0527	0.3261 ± 0.0296	0.8590 ± 0.0062		
Client 2 accuracy	0.9266 ± 0.0125	0.9060 ± 0.0157	0.9035 ± 0.0162	0.9357 ± 0.0037		
Client 3 accuracy	0.6724 ± 0.0467	0.3279 ± 0.0353	0.3612 ± 0.0814	0.7896 ± 0.0109		
Client 4 accuracy	0.7979 ± 0.0448	0.6858 ± 0.0105	0.6654 ± 0.0121	0.8098 ± 0.0112		
Client 5 accuracy	0.6468 ± 0.0248	0.4548 ± 0.0655	0.4234 ± 0.0387	0.7426 ± 0.0257		

Figure 5: In this experiment with Fashion MNIST, a simple MLP with dropout were employed.

 Table 7: ResNet-18 on CIFAR-10 with label shift across 5 clients. For fair comparison, we run 5,000 iterations for our method and Upper Bound, while 10000 for FedAvg and FedBN.

1852 1853	CIFAR-10	Our IW-ERM	FedAvg	FedBN	Upper Bound		
1854	Avg. accuracy	0.5640 ± 0.0241	0.4515 ± 0.0148	0.4263 ± 0.0975	0.5790 ± 0.0103		
1855	Client 1 accuracy	0.6410 ± 0.0924	0.5405 ± 0.1845	$0.5321 + 0.0620$	0.7462 ± 0.0339		
	Client 2 accuracy	0.8434 ± 0.0359	$0.3753 + 0.0828$	0.4656 ± 0.2158	0.7509 ± 0.0534		
	Client 3 accuracy	0.4591 ± 0.1131	0.3973 ± 0.1333	0.2838 ± 0.1055	0.5845 ± 0.0854		
	Client 4 accuracy	0.4751 ± 0.1241	0.5007 ± 0.1303	0.5256 ± 0.1932	0.3507 ± 0.0578		
	Client 5 accuracy	0.4013 ± 0.0430	0.4429 ± 0.1195	0.5603 ± 0.1581	0.4627 ± 0.0456		

 Figure 6: The average, best-client, and worst-client accuracy, along with their standard deviations, are derived from Table $\overline{6}$. Our method exhibits the lowest standard deviation, showcasing the most robust accuracy amongst the compared methods.

 Figure 7: The average, best-client, and worst-client accuracy, along with their standard deviations, are derived from Table [7.](#page-20-0)

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Table 8: Label distribution on Fasion MNIST with 5 clients, with the majority of classes possessing a limited number of training and test images across each client.

Table 9: Label distribution on CIFAR-10 with 5 clients, with the majority of classes possessing a limited number of training and test images across each client.

							Class				
		0		2	3	4	5	6		8	9
Client 1	Train	34	34	34	34	34	5862	34	34	34	34
	Test	977		5		5					
Client 2	Train	34	34	34	34	34	34	5862	34	34	34
	Test	5	977		5	5	5				
Client 3	Train	34	34	34	34	34	34	34	5862	34	34
	Test	5	5	977	5	5	5	5			
Client 4	Train	34	34	34	34	34	34	34	34	5862	34
	Test		5	5	977	5	5				
Client 5	Train	34	34	34	34	34	34	34	34	34	5862
	Test	5	5	5	5	977					

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1948 1949 1950 Table 10: Label distribution on CIFAR-10 with 100 clients, wherein groups of 10 clients share the same distribution and ratios. The majority of classes possess a limited quantity of training and test images on each client.

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