
Supplementary Material: Optimizing Solution-Samplers for Combinatorial Problems: The Landscape of Policy-Gradient Methods

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Abstract

1 Deep Neural Networks and Reinforcement Learning methods have empirically
2 shown great promise in tackling challenging combinatorial problems. In those
3 methods a deep neural network is used as a solution generator which is then trained
4 by gradient-based methods (e.g., policy gradient) to successively obtain better
5 solution distributions. In this work we introduce a novel theoretical framework for
6 analyzing the effectiveness of such methods. We ask whether there exist generative
7 models that (i) are expressive enough to generate approximately optimal solutions;
8 (ii) have a tractable, i.e. polynomial in the size of the input, number of parameters;
9 (iii) their optimization landscape is benign in the sense that it does not contain
10 sub-optimal stationary points. Our main contribution is a positive answer to this
11 question. Our result holds for a broad class of combinatorial problems including
12 Max- and Min-Cut, Max- k -CSP, Maximum-Weight-Bipartite-Matching, and the
13 Traveling Salesman Problem. As a byproduct of our analysis we introduce a novel
14 regularization process over vanilla gradient descent and provide theoretical and
15 experimental evidence that it helps address vanishing-gradient issues and escape
16 bad stationary points.

17 1 Introduction

18 Gradient descent has proven remarkably effective for diverse optimization problems in neural net-
19 works. From the early days of neural networks, this has motivated their use for combinatorial
20 optimization [HT85, Smi99, VFJ15, BPL⁺16]. More recently, an approach by [BPL⁺16], where a
21 neural network is used to generate (sample) solutions for the combinatorial problem. The parameters
22 of the neural network thus parameterize the space of distributions. This allows one to perform gradient
23 steps in this distribution space. In several interesting settings, including the Traveling Salesman Prob-
24 lem, they have shown that this approach works remarkably well. Given the widespread application but
25 also the notorious difficulty of combinatorial optimization [GLS12, PS98, S⁺03, Sch05, CLS⁺95],
26 approaches that provide a more general solution framework are appealing.

27 This is the point of departure of this paper. We investigate whether gradient descent can succeed in a
28 general setting that encompasses the problems studied in [BPL⁺16]. This requires a parameterization
29 of distributions over solutions with a “nice” optimization landscape (intuitively, that gradient descent
30 does not get stuck in local minima or points of vanishing gradient) and that has a polynomial
31 number of parameters. Satisfying both requirements simultaneously is non-trivial. As we show
32 precisely below, a simple lifting to the exponential-size probability simplex on all solutions guarantees
33 convexity; and, on the other hand, *compressed* parameterizations with “bad” optimization landscapes
34 are also easy to come by (we give a natural example for Max-Cut in Remark 1). Hence, we seek to

understand the parametric complexity of gradient-based methods, i.e., how many parameters suffice for a benign optimization landscape in the sense that it does not contain “bad” stationary points.

We thus theoretically investigate whether there exist solution generators with a tractable number of parameters that are also efficiently optimizable, i.e., gradient descent requires a small number of steps to reach a near-optimal solution. We provide a positive answer under general assumptions and specialize our results for several classes of *hard and easy* combinatorial optimization problems, including Max-Cut and Min-Cut, Max- k -CSP, Maximum-Weighted-Bipartite-Matching and Traveling Salesman. We remark that a key difference between (computationally) easy and hard problems is not the ability to find a compressed and efficiently optimizable generative model but rather the ability to efficiently draw samples from the parameterized distributions.

1.1 Our Framework

We introduce a theoretical framework for analyzing the effectiveness of gradient-based methods on the optimization of solution generators in combinatorial optimization, inspired by [BPL⁺16].

Let \mathcal{I} be a collection of instances of a combinatorial problem with common solution space S and $L(\cdot; I) : S \rightarrow \mathbb{R}$ be the cost function associated with an instance $I \in \mathcal{I}$, i.e., $L(s; I)$ is the cost of solution s given the instance I . For example, for the Max-Cut problem the collection of instances \mathcal{I} corresponds to all graphs with n nodes, the solution space S consists of all subsets of nodes, and the loss $L(s; I)$ is equal to (minus) the weight of the cut $(s, [n] \setminus s)$ corresponding to the subset of nodes $s \in S$ (our goal is to minimize L).

Definition 1 (Solution Cost Oracle). *For a given instance I we assume that we have access to an oracle $\mathcal{O}(\cdot; I)$ to the cost of any given solution $s \in S$, i.e., $\mathcal{O}(s; I) = L(s; I)$.*

The above oracle is standard in combinatorial optimization and query-efficient algorithms are provided for various problems [RSW17, GPRW19, LSZ21, AEG⁺22, PRW22]. We remark that the goal of this work is not to design algorithms that solve combinatorial problems using access to the solution cost oracle (as the aforementioned works do). This paper focuses on landscape design: the algorithm is **fixed**, namely (stochastic) gradient descent; the question is how to design a generative model that has a small number of parameters and the induced optimization landscape allows gradient-based methods to converge to the optimal solution without getting trapped at local minima or vanishing gradient points.

Let \mathcal{R} be some prior distribution over the instance space \mathcal{I} and \mathcal{W} be the space of parameters of the model. We now define the class of solution generators. The solution generator $p(w)$ with **parameter** $w \in \mathcal{W}$ takes as **input** an instance I and **generates a random solution** s in S . To distinguish between the output, the input, and the parameter of the solution generator, we use the notation $p(\cdot; I; w)$ to denote the distribution over solutions and $p(s; I; w)$ to denote the probability of an individual solution $s \in S$. We denote by $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$ the above parametric class of solution generators. For some parameter w , the loss corresponding to the solutions sampled by $p(\cdot; I; w)$ is equal to

$$\mathcal{L}(w) = \mathbf{E}_{I \sim \mathcal{R}} [\mathcal{L}_I(w)], \quad \mathcal{L}_I(w) = \mathbf{E}_{s \sim p(\cdot; I; w)} [L(s; I)]. \quad (1)$$

Our goal is to optimize the parameter $w \in \mathcal{W}$ in order to find a sampler $p(\cdot; I; w)$ whose loss $\mathcal{L}(w)$ is close to the expected optimal value opt :

$$\text{opt} = \mathbf{E}_{I \sim \mathcal{R}} \left[\min_{s \in S} L(s; I) \right]. \quad (2)$$

The policy gradient method [Kak01] expresses the gradient of \mathcal{L} as follows

$$\nabla_w \mathcal{L}(w) = \mathbf{E}_{I \sim \mathcal{R}} \mathbf{E}_{s \sim p(\cdot; I; w)} [L(s; I) \nabla_w \log p(s; I; w)],$$

and updates the parameter w using the gradient descent update. Observe that a (stochastic) policy gradient update can be implemented using only access to a solution cost oracle of Definition 1.

Solution Generators. In [BPL⁺16] the authors used neural networks as parametric solution generators for the TSP problem. They provided empirical evidence that optimizing the parameters of the neural network using the policy gradient method results to samplers that generate very good

solutions for (Euclidean) TSP instances. Parameterizing the solution generators using neural networks essentially *compresses* the description of distributions over solutions (the full parameterization would require assigning a parameter to every solution-instance pair (s, I)). Since for most combinatorial problems the size of the solution space is exponentially large (compared to the description of the instance), it is crucial that for such methods to succeed the parameterization must be *compressed* in the sense that the description of the parameter space \mathcal{W} is polynomial in the size of the description of the instance family \mathcal{I} . Apart from having a tractable number of parameters, it is important that the *optimization objective* corresponding to the parametric class \mathcal{P} can provably be optimized using some first-order method in polynomial (in the size of the input) iterations.

We collect these desiderata in the following definition. We denote by $[\mathcal{I}]$ the description size of \mathcal{I} , i.e., the number of bits required to identify any element of \mathcal{I} . For instance, if \mathcal{I} is the space of unweighted graphs with at most n nodes, $[\mathcal{I}] = O(n^2)$.

Definition 2 (Complete, Compressed and Efficiently Optimizable Solution Generator). *Fix a prior \mathcal{R} over \mathcal{I} , a family of solution generators $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$, a loss function \mathcal{L} as in Equation (1) and some $\epsilon > 0$.*

1. We say that \mathcal{P} is **complete** if there exists some $\bar{w} \in \mathcal{W}$ such that $\mathcal{L}(\bar{w}) \leq \text{opt} + \epsilon$, where opt is defined in (2).
2. We say that \mathcal{P} is **compressed** if the description size of the parameter space \mathcal{W} is polynomial in $[\mathcal{I}]$ and in $\log(1/\epsilon)$.
3. We say that \mathcal{P} is **efficiently optimizable** if there exists a first-order method applied on the objective \mathcal{L} such that after $T = \text{poly}([\mathcal{W}], 1/\epsilon)$ many updates on the parameter vectors, finds an (at most) ϵ -sub-optimal vector \hat{w} , i.e., $\mathcal{L}(\hat{w}) \leq \mathcal{L}(\bar{w}) + \epsilon$.

Remark 1. We remark that constructing parametric families that are complete and compressed, complete and efficiently optimizable, or compressed and efficiently optimizable (i.e., satisfying any pair of assumptions of [Question 1](#) but not all 3) is usually a much easier task. Consider, for example, the Max-Cut problem on a fixed (unweighted) graph with n nodes. Note that \mathcal{I} has description size $O(n^2)$. Solutions of the Max-Cut for a graph with n nodes are represented by vertices on the binary hypercube $\{\pm 1\}^n$ (coordinate i dictates the side of the cut that we put node i). One may consider the full parameterization of all distributions over the hypercube. It is not hard to see that this is a **complete and efficiently optimizable** family (the optimization landscape corresponds to optimizing a linear objective). However, it is **not compressed**, since it requires 2^n parameters. On the other extreme, considering a product distribution over coordinates, i.e., we set the value of node i to be $+1$ with probability p_i and -1 with $1 - p_i$ gives a **complete and compressed** family. However, as we show in [Appendix B](#), the landscape of this compressed parameterization suffers from highly sub-optimal local minima and therefore, it is **not efficiently optimizable**.

Therefore, in this work we investigate whether it is possible to have all 3 desiderata of [Definition 2](#) at the same time.

Question 1. *Are there complete, compressed, and efficiently optimizable solution generators (i.e., satisfying [Definition 2](#)) for challenging combinatorial tasks?*

1.2 Our Results

Our Contributions. Before we present our results formally, we summarize the contributions of this work.

- Our main contribution is a positive answer ([Theorem 1](#)) to [Question 1](#) under general assumptions that capture many combinatorial tasks. We identify a set of conditions (see [Assumption 1](#)) that allow us to design a family of solution generators that are complete, compressed and efficiently optimizable.
- The conditions are motivated by obstacles that are important for any approach of this nature. This includes solutions that escape to infinity, and also parts of the landscape with vanishing gradient. See the discussion in [Section 3](#) and [Figure 1](#).
- We specialize our framework to several important combinatorial problems, some of which are NP-hard, and others tractable: Max-Cut, Min-Cut, Max- k -CSP, Maximum-Weight-Bipartite-Matching, and the Traveling Salesman Problem.

- Finally, we investigate experimentally the effect of the entropy regularizer and the fast/slow mixture scheme that we introduced (see [Section 3](#)) and provide evidence that it leads to better solution generators.

We begin with the formal presentation of our assumptions on the feature mappings of the instances and solutions and on the structure of cost function of the combinatorial problem.

Assumption 1 (Structured Feature Mappings). *Let S be the solution space and \mathcal{I} be the instance space. There exist feature mappings $\psi_S : S \rightarrow X$, for the solutions, and, $\psi_{\mathcal{I}} : \mathcal{I} \rightarrow Z$, for the instances, where X, Z are Euclidean vector spaces of dimension n_X and n_Z , such that*

1. (Bounded Feature Spaces) *The feature and instance mappings are bounded, i.e., there exist $D_S, D_{\mathcal{I}} > 0$ such that $\|\psi_S(s)\|_2 \leq D_S$, for all $s \in S$ and $\|\psi_{\mathcal{I}}(I)\|_2 \leq D_{\mathcal{I}}$, for all $I \in \mathcal{I}$.*
2. (Bilinear Cost Oracle) *The cost of a solution s under instance I can be expressed as a bilinear function of the corresponding feature vector $\psi_S(s)$ and instance vector $\psi_{\mathcal{I}}(I)$, i.e., the solution oracle can be expressed as $\mathcal{O}(s, I) = \psi_{\mathcal{I}}(I)^\top M \psi_S(s)$ for any $s \in S, I \in \mathcal{I}$, for some matrix M with $\|M\|_F \leq C$.*
3. (Variance Preserving Features) *There exists $\alpha > 0$ such that $\text{Var}_{s \sim U(S)}[v \cdot \psi_S(s)] \geq \alpha \|v\|_2^2$ for any $v \in X$, where $U(S)$ is the uniform distribution over the solution space S .*
4. (Bounded Dimensions/Diameters) *The feature dimensions n_X, n_Z , and the diameter bounds $D_S, D_{\mathcal{I}}, C$ are bounded above by a polynomial of the description size of the instance space \mathcal{I} . The variance lower bound α is bounded below by $1/\text{poly}(|\mathcal{I}|)$.*

Remark 2 (Boundedness and Bilinear Cost Assumptions). *We remark that Items 1, 4 are simply boundedness assumptions for the corresponding feature mappings and usually follow easily assuming that we consider reasonable feature mappings. At a high-level, the assumption that the solution is a bilinear function of the solution and instance features (Item 2) prescribes that “good” feature mappings should enable a simple (i.e., bilinear) expression for the cost function. In the sequel we see that this is satisfied by natural feature mappings for important classes of combinatorial problems.*

Remark 3 (Variance Preservation Assumption). *In Item 3 (variance preservation) we require that the solution feature mapping has variance along every direction, i.e., the feature vectors corresponding to the solutions must be “spread-out” when the underlying solution generator is the uniform distribution. As we show, this assumption is crucial so that the gradients of the resulting optimization objective are not-vanishing, allowing for its efficient optimization.*

We mention that various important combinatorial problems satisfy [Assumption 1](#). For instance, [Assumption 1](#) is satisfied by Max-Cut, Min-Cut, Max- k -CSP, Maximum-Weight-Bipartite-Matching and Traveling Salesman. We refer the reader to the upcoming [Section 2](#) for an explicit description of the structured feature mappings for these problems. Having discussed [Assumption 1](#), we are ready to state our main abstract result which resolves [Question 1](#).

Theorem 1. *Consider a combinatorial problem with instance space \mathcal{I} that satisfies [Assumption 1](#). For any prior \mathcal{R} over \mathcal{I} and $\epsilon > 0$, there exists a family of solution generators $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$ with parameter space \mathcal{W} that is complete, compressed and, efficiently optimizable.*

A sketch behind the design of the family \mathcal{P} can be found in [Section 3](#) and [Section 4](#).

Remark 4 (Computational Barriers in Sampling). *We note that the families of generative models (a.k.a., solution generators) that we provide have polynomial parameter complexity and are optimizable in a small number of steps using gradient-based methods. Hence, in a small number of iterations, gradient-based methods converge to distributions whose mass is concentrated on nearly optimal solutions. This holds, as we show, even for challenging (NP-hard) combinatorial problems. Our results do not, however, prove $P = NP$, as it may be computationally hard to sample from our generative models. We remark that while such approaches are in theory hard, such models seem to perform remarkably well experimentally where sampling is based on Langevin dynamics techniques [[SE20](#), [SSDK⁺20](#)]. Though as our theory predicts, and simulations support, landscape problems seem to be a direct impediment even to obtain good approximate solutions.*

Remark 5 (Neural Networks as Solution Samplers). *A natural question would be whether our results can be extended to the case where neural networks are (efficient) solution samplers, as in [[BPL⁺16](#)]. Unfortunately, a benign landscape result for neural network solution generators most likely cannot*

183 *exist. It is well-known that end-to-end theoretical guarantees for training neural networks are out of*
 184 *reach since the corresponding optimization tasks are provably computationally intractable, see, e.g.,*
 185 *[CGKM22] and the references therein.*

186 1.3 Related Work

187 Tackling combinatorial optimization problems constitutes one of the most fundamental tasks of
 188 theoretical computer science [GLS12, PS98, S⁺03, Sch05, CLS⁺95] and various approaches have
 189 been studied for these problems such as local search methods, branch-and-bound algorithms and
 190 meta-heuristics such as genetic algorithms and simulated annealing. Starting from the seminal
 191 work of [HT85], researchers apply neural networks [Smi99, VFJ15, BPL⁺16] to solve combina-
 192 torial optimization tasks. In particular, researchers have explored the power of machine learning,
 193 reinforcement learning and deep learning methods for solving combinatorial optimization prob-
 194 lems [BPL⁺16, YW20, LZ09, DCL⁺18, BLP21, MSIB21, NOST18, SHM⁺16, MKS⁺13, SSS⁺17,
 195 ER18, KVHW18, ZCH⁺20, CCK⁺21, MGH⁺19, GCF⁺19, KLMS19].

196 The training phase of the neural network approaches can be distinguished in two directions. The
 197 first direction deals with supervised training [SLB⁺18, JLB19, GCF⁺19]; however, supervised
 198 training exhibits generalization issues [YGS20]. The second direction deals with unsupervised train-
 199 ing [MSIB21, BPL⁺16, KDZ⁺17, YP19, CT19, YBV19, KCK⁺20, KCY⁺21, DAT20, NJS⁺20,
 200 TRWG21, AMW18, KL20, SBK22]. Our paper builds on the framework of the influential experi-
 201 mental work of [BPL⁺16] to tackle combinatorial optimization problems such as TSP using neural
 202 networks and RL. [KP⁺21] uses an entropy maximization scheme in order to generate diversified
 203 candidate solutions. This experimental heuristic is quite close to our theoretical idea for entropy
 204 regularization. In our work, entropy regularization allows us to design quasar-convex landscapes
 205 and the fast/slow mixing scheme to obtain diversification of solutions. Among other related applied
 206 works, [KCK⁺20, KPP22] study the use of Transformer architectures combined with the Reinforce
 207 algorithm employing symmetries (i.e., the existence of multiple optimal solutions of a CO problem)
 208 improving the generalization capability of Deep RL NCO and [MLC⁺21] studies Transformer archi-
 209 tectures and aims to learn improvement heuristics for routing problems using RL. Finally, this
 210 work lies in the area of providing theoretical evidence and insight for the practical success of deep
 211 learning and neural networks (see [AS18, MSS19, AS20, CNP20, ABAB⁺21, MYSSS21, BEG⁺22,
 212 GKKZ22, ABA22, AAM22, CCL⁺22, ABAM23] for a small sample of this line of research).

213 2 Combinatorial Applications

214 We now consider concrete combinatorial problems and show that there exist appropriate and natural
 215 feature mappings for the solutions and instances that satisfy [Assumption 1](#); so [Theorem 1](#) is applicable
 216 for any such combinatorial task. For a more detailed treatment, we refer to [Appendix G](#).

217 **Min-Cut and Max-Cut.** Min-Cut (resp. Max-Cut) are central graph combinatorial problems where
 218 the task is to split the nodes of the graph in two subsets so that the number of edges from one subset
 219 to the other (edges of the cut) is minimized (resp. maximized). Given a graph G with n nodes
 220 represented by its Laplacian matrix $L_G = D - A$, where D is the diagonal degree matrix and A
 221 is the adjacency matrix of the graph, the goal in the Min-Cut (resp. Max-Cut) problem is to find a
 222 solution vector $s \in \{\pm 1\}^n$ so that $s^\top L_G s / 4$ is minimized (resp. maximized).

223 We first show that there exist natural feature mappings so that the cost of every solution s under
 224 any instance/graph G is a bilinear function of the feature vectors, see Item 2 of [Assumption 1](#).
 225 We consider the correlation-based feature mapping $\psi_S(s) = (ss^\top)^b \in \mathbb{R}^{n^2}$, where by $(\cdot)^b$ we
 226 denote the vectorization/flattening operation and the Laplacian for the instance (graph), $\psi_{\mathcal{I}}(G) =$
 227 $(L_G)^b \in \mathbb{R}^{n^2}$. Then simply setting the matrix M to be the identity $I \in \mathbb{R}^{n^2 \times n^2}$ the cost of any
 228 solution s can be expressed as the bilinear function $\psi_{\mathcal{I}}(G)^\top M \psi_S(s) = (L_G^b)^\top (ss^\top)^b = s^\top L_G s$.
 229 We observe that for (unweighted) graphs with n nodes the description size of the family of all
 230 instances \mathcal{I} is roughly $O(n^2)$, and therefore the dimensions of the feature mappings $\psi_S, \psi_{\mathcal{I}}$ are
 231 clearly polynomial in the description size of \mathcal{I} . Moreover, considering unweighted graphs, it holds
 232 that $\|\psi_{\mathcal{I}}(G)\|_2, \|\psi_S(s)\|_2, \|M\|_F \leq \text{poly}(n)$. Therefore, the constants $D_S, D_{\mathcal{I}}, C$ are polynomial in
 233 the description size of the instance family.

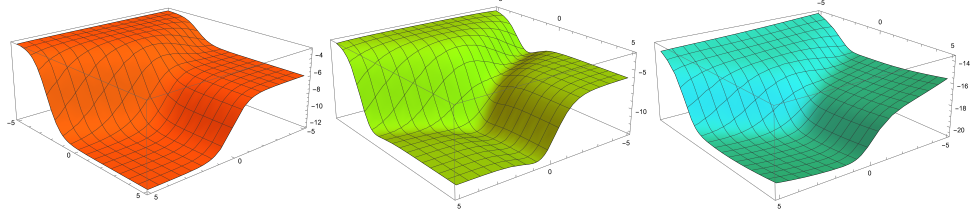


Figure 1: In the **left** plot, we show the landscape of the “vanilla” objective of Eq.(1) for the feature domain $X = \{(1, 0), (2, 2), (0, 2)\}$ and linear cost oracle $c \cdot x$ for $c = (-3, -3)$. We see that the “vanilla” objective is minimized at the direction of $-c$, i.e., along the direction $\tau(1, 1)$ for $\tau \rightarrow +\infty$. We observe the two issues described in Section 3, i.e., that the true minimizer is a point at infinity, and that gradients vanish so gradient descent may get trapped in sub-optimal solutions, (e.g., in the upper-right corner if initialized in the top corner). In the **middle** plot, we show the landscape of the entropy-regularized objective of Eq.(3) that makes the minimizer finite and brings it closer to the origin. Observe that even if a gradient iteration is initialized in the top corner it will eventually converge to the minimizer; however the rate of convergence may be very slow. The **right** plot corresponds to the loss objective where we combine a mixture of exponential families as solution generator, as in Eq.(5), and the entropy regularization approach. We observe that we are able to obtain a benign (quasar-convex) landscape via the entropy regularization while the mixture-generator guarantees non-vanishing gradients.

It remains to show that our solution feature mapping satisfies the variance preservation assumption, i.e., Item 3 in Assumption 1. A uniformly random solution vector $s \in \{\pm 1\}^n$ is sampled by setting each $s_i = 1$ with probability $1/2$ independently. In that case, we have $\mathbf{E}[v \cdot x] = 0$ and therefore $\text{Var}(v \cdot x) = \mathbf{E}[(v \cdot x)^2] = \sum_{i \neq j} v_i v_j \mathbf{E}[x_i x_j] = \sum_i v_i^2 = \|v\|_2^2$, since, by the independence of x_i, x_j , the cross-terms of the sum vanish. We observe that the same hold true for the Max-Cut problem and therefore, structured feature mappings exist for Max-Cut as well (where $L(s; G) = -s^\top L_G s$). We shortly mention that there also exist structured feature mappings for Max- k -CSP. We refer to Theorem 4 for further details.

Remark 6 (Partial Instance Information/Instance Context). We remark that Assumption 1 allows for the “instance” I to only contain partial information about the actual cost function. For example, consider the setting where each sampled instance is an unweighted graph G but the cost oracle takes the form $\mathcal{O}(G, s) = (L_G)^p M(ss^\top)^p$ for a matrix $M_{ij} = a_i$ when $i = j$ and $M_{ij} = 0$ otherwise. This cost function models having a **unknown weight function**, i.e., the weight of edge i of G is a_i if edge i exists in the observed instance G , on the edges of the observed unweighted graph G , that the algorithm has to learn in order to be able to find the minimum or maximum cut. For simplicity, in what follows, we will continue referring to I as the instance even though it may only contain partial information about the cost function of the underlying combinatorial problem.

Maximum-Weight-Bipartite-Matching and TSP. The Maximum-Weight-Bipartite-Matching (MWBP) problem is another graph problem that, given a bipartite graph G with n nodes and m edges, asks for the maximum-weight matching. The feature vector corresponding to a matching can be represented as a binary matrix $R \in \{0, 1\}^{n \times n}$ with $\sum_j R_{ij} = 1$ for all i and $\sum_i R_{ij} = 1$ for all j , i.e., R is a permutation matrix. Therefore, for a candidate matching s , we set $\psi_S(s)$ to be the matrix R defined above. Moreover, the feature vector of the graph is the (negative flattened) adjacency matrix E^b . The cost oracle is then $L(R; E) = \sum_{ij} E_{ij} M_{ij} R_{ij}$ perhaps for an unknown weight matrix M_{ij} (see Remark 6). For the Traveling Salesman Problem (TSP) the feature vector is again a matrix R with the additional constraint that R has to represent a single cycle (a tour over all cities). The cost function for TSP is again $L(R; E) = \sum_{ij} E_{ij} M_{ij} R_{ij}$. One can check that those representations of the instance and solution satisfy the assumptions of Items 1 and 4. Showing that the variance of those representations has a polynomial lower bound is more subtle and we refer the reader to the Supplementary Material.

We shortly mention that the solution generators for Min-Cut and Maximum-Weight-Bipartite-Matching are also efficiently samplable.

3 Optimization Landscape

Exponential Families as Solution Generators. A natural candidate to construct our family of solution generators is to consider the distribution that assigns to each solution $s \in S$ and instance $I \in \mathcal{I}$ mass proportional to its score $\exp(-\tau L(s; I)) = \exp(-\tau \psi_{\mathcal{I}}(I)^\top M \psi_S(s)) = \exp(-\tau z^\top M x)$ for some “temperature” parameter τ , where $\psi_{\mathcal{I}}$ and ψ_S are the feature mappings promised to exist due to [Assumption 1](#), $z = \psi_{\mathcal{I}}(I)$, and, $x = \psi_S(s)$. Note that as long as $\tau \rightarrow +\infty$, this distribution tends to concentrate on solutions that achieve small loss.

Remark 7. *To construct the above solution sampler one could artificially query specific solutions to the cost oracle of [Definition 1](#) and try to learn the cost matrix M . However, we remark that our goal (see [Definition 2](#)) is to show that we can train a parametric family via gradient-based methods so that it generates (approximately) optimal solutions and not to simply learn the cost matrix M via some other method and then use it to generate good solutions.*

Obstacle I: Minimizers at Infinity. One could naturally consider the parametric family $\phi(x; z; W) \propto \exp(z^\top W x)$ (note that with $W = -\tau M$, we recover the distribution of the previous paragraph) and try to perform gradient-based methods on the loss (recall that $L(x; z) = z^\top M x$)¹

$$\mathcal{L}(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim \phi(\cdot; z; W)} [z^\top M x]. \quad (1)$$

The question is whether gradient updates on the parameter W eventually converge to a matrix \bar{W} whose associated distribution $\phi(\bar{W})$ generates near-optimal solutions (note that the matrix $-\tau M$ with $\tau \rightarrow +\infty$ is such a solution). After computing the gradient of \mathcal{L} , we observe that

$$\nabla_W \mathcal{L}(W) \cdot M = \mathbf{Var}_{z \sim \mathcal{R}, x \sim \phi(\cdot; z; W)} [z^\top M x] \geq 0,$$

where the inner product between two matrices $A \cdot B$ is the trace $\text{Tr}(A^\top B) = \sum_{i,j} A_{ij} B_{ij}$. This means that the gradient field of \mathcal{L} always has a contribution to the direction of M . Nevertheless the actual minimizer is at infinity, i.e., it corresponds to the point $\bar{W} = -\tau M$ when $\tau \rightarrow +\infty$. While the correlation with the optimal point is positive (which is encouraging), having such contribution to this direction is not a sufficient condition for actually reaching \bar{W} . The objective has vanishing gradients at infinity and gradient descent may get trapped in sub-optimal stationary points, see the left plot in [Figure 1](#).

Solution I: Quasar Convexity via Entropy Regularization. Our plan is to try and make the objective landscape more benign by adding an entropy-regularizer. Instead of trying to make the objective convex (which may be too much to ask in the first place) *we are able obtain a much better landscape with a finite global minimizer and a gradient field that guides gradient descent to the minimizer.* Those properties are described by the so-called class of “quasar-convex” functions. Quasar convexity (or weak quasi-convexity [[HMR16](#)]) is a well-studied notion in optimization [[HMR16](#), [HSS20](#), [LV16](#), [ZMB⁺17](#), [HLSS15](#)] and can be considered as a high-dimensional generalization of unimodality.

Definition 3 (Quasar Convexity [[HMR16](#), [HSS20](#)]). *Let $\gamma \in (0, 1]$ and let \bar{x} be a minimizer of the differentiable function $f : \mathbb{R}^n \rightarrow \mathbb{R}$. The function f is γ -quasar-convex with respect to \bar{x} on a domain $D \subseteq \mathbb{R}^n$ if for all $x \in D$, $\nabla f(x) \cdot (x - \bar{x}) \geq \gamma(f(x) - f(\bar{x}))$.*

In the above definition, notice that the main property that we need to establish is that the gradient field of our objective correlates positively with the direction $W - \bar{W}$, where \bar{W} is its minimizer. We denote by $H : \mathcal{W} \rightarrow \mathbb{R}$ the negative entropy of $\phi(W)$, i.e.,

$$H(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim \phi(\cdot; z; W)} [\log \phi(x; z; W)], \quad (2)$$

and consider the *regularized* objective

$$\mathcal{L}_\lambda(W) = \mathcal{L}(W) + \lambda H(W), \quad (3)$$

¹We note that we overload the notation and assume that our distributions generate directly the featurizations z (resp. x) of I (resp. s).

for some $\lambda > 0$. We show (follows from [Lemma 4](#)) that the gradient-field of the regularized objective indeed “points” towards a finite minimizer (the matrix $\bar{W} = -M/\lambda$):

$$\begin{aligned} \nabla_W \mathcal{L}_\lambda(W) \cdot (W + M/\lambda) = \\ \mathbf{Var}[z^\top (W + M/\lambda)x] \geq 0, \end{aligned} \quad (4)$$

where the randomness is over $z \sim \mathcal{R}, x \sim \phi(\cdot; z; W)$. Observe that now the minimizer of \mathcal{L}_λ is the point $-M/\lambda$, which for $\lambda = \text{poly}(\epsilon, \alpha, 1/C, 1/D_S, 1/D_{\mathcal{I}})$ (these are the parameters of [Assumption 1](#)) is promised to yield a solution sampler that generates ϵ -sub-optimal solutions (see also [Proposition 2](#) and [Appendix C](#)). Having the property of [Equation \(4\)](#) suffices for showing that a gradient descent iteration (with an appropriately small step-size) will *eventually* converge to the minimizer.

Obstacle II: Vanishing Gradients. While we have established that the gradient field of the regularized objective “points” towards the right direction, the regularized objective still suffers from vanishing gradients, see the middle plot in [Figure 1](#). In other words, γ in the definition of quasar convexity ([Definition 3](#)) may be exponentially small, as it is proportional to the variance of the random variable $z^\top (W + M/\lambda)x$, see [Equation \(4\)](#). As we see in the middle plot of [Figure 1](#), the main issue is the vanishing gradient when W gets closer to the minimizer $-M/\lambda$ (towards the front-corner). For simplicity, consider the variance along the direction of W , i.e., $\mathbf{Var}[z^\top Wx]$ and recall that x is generated by the density $\exp(z^\top Wx)/(\sum_{x \in X} \exp(z^\top Wx))$. When $\|W\|_2 \rightarrow +\infty$ we observe that the value $z^\top Wx$ concentrates exponentially fast to $\max_{x \in X} z^\top Wx$ (think of the convergence of the soft-max to the max function). Therefore, the variance $\mathbf{Var}[z^\top Wx]$ may vanish exponentially fast making the convergence of gradient descent slow.

Solution II: Non-Vanishing Gradients via Fast/Slow Mixture Generators. We propose a fix to the vanishing gradients issue by using a mixture of exponential families as a solution generator. We define the family of solution generators $\mathcal{P} = \{p(W) : W \in \mathcal{W}\}$ to be

$$\mathcal{P} = \{(1 - \beta^*)\phi(W) + \beta^*\phi(\rho^*W) : W \in \mathcal{W}\}, \quad (5)$$

for a (fixed) mixing parameter β^* and a (fixed) temperature parameter ρ^* . The main idea is to have the first component of the mixture to converge fast to the optimal solution (to $-M/\lambda$) while the second “slow” component that has parameter ρ^*W stays closer to the uniform distribution over solutions that guarantees non-trivial variance (and therefore non-vanishing gradients).

More precisely, taking ρ^* to be sufficiently small, the distribution $\phi(\rho^*W)$ is *almost uniform* over the solution space $\psi_S(S)$. Therefore, in [Equation \(4\)](#), the almost uniform distribution component of the mixture will add to the variance and allow us to show a lower bound. This is where Item 3 of [Assumption 1](#) comes into play and gives us the desired non-trivial variance lower bound under the uniform distribution. We view this fast/slow mixture technique as an interesting insight of our work: we use the “fast” component (the one with parameter W) to actually reach the optimal solution $-M/\lambda$ and we use the “slow” component (the one with parameter ρ^*W that essentially generates random solutions) to preserve a non-trivial variance lower bound during optimization.

4 Complete, Compressed and Efficiently Optimizable Solution Generators

In this section, we discuss the main results that imply [Theorem 1](#): the family \mathcal{P} of [Equation \(5\)](#) is complete, compressed and efficiently optimizable (for some choice of β^*, ρ^* and \mathcal{W}).

Completeness. First, we show that the family of solution generators of [Equation \(5\)](#) is complete. For the proof, we refer to [Proposition 2](#) in [Appendix C](#). At a high-level, we to pick β^*, ρ^* to be of order $\text{poly}(\epsilon, \alpha, 1/C, 1/D_S, 1/D_{\mathcal{I}})$. This yields that the matrix $\bar{W} = -M/\lambda$ is such that $\mathcal{L}(\bar{W}) \leq \text{opt} + \epsilon$, where M is the matrix of Item 2 in [Assumption 1](#) and λ is $\text{poly}(\epsilon/[\mathcal{I}])$. To give some intuition about this choice of matrix, let us see how $\mathcal{L}(\bar{W})$ behaves. By definition, we have that

$$\mathcal{L}(\bar{W}) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; \bar{W})} [z^\top Mx],$$

where the distribution p belongs to the family of [Equation \(5\)](#), i.e., $p(\bar{W}) = (1 - \beta^*)\phi(\bar{W}) + \beta^*\phi(\rho^*\bar{W})$. Since the mixing weight β^* is small, we have that $p(\bar{W})$ is approximately equal to

350 $\phi(\bar{W})$. This means that our solution generator draws samples from the distribution whose mass at x
 351 given instance z is proportional to $\exp(-z^\top Mx/\lambda)$ and, since $\lambda > 0$ is very small, the distribution
 352 concentrates to solutions x that tend to minimize the objective $z^\top Mx$. This is the reason why
 353 $\bar{W} = -M/\lambda$ is close to opt in the sense that $\mathcal{L}(\bar{W}) \leq \text{opt} + \epsilon$.

354 **Compression.** As a second step, we show (in [Proposition 3](#), see [Appendix D](#)) that \mathcal{P} is a compressed
 355 family of solution generators. This result follows immediately from the structure of Equation (5)
 356 (observe that W has $n_X n_Z$ parameters) and the boundedness of $\bar{W} = -M/\lambda$.

357 **Efficiently Optimizable.** The proof of this result essentially corresponds to the discussion provided
 358 in [Section 3](#). Our main structural result shows that the landscape of the regularized objective with the
 359 fast/slow mixture solution-generator is quasr convex. More precisely, we consider the following
 360 objective:

$$\mathcal{L}_\lambda(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; W)} [z^\top Mx] + \lambda R(W), \quad (1)$$

361 where $p(W)$ belongs in the family \mathcal{P} of Equation (5) and R is a weighted sum of two negative entropy
 362 regularizers (to be in accordance with the mixture structure of \mathcal{P}), i.e., $R(W) = (1 - \beta^*)H(W) +$
 363 $\beta^*/\rho^* H(\rho^* W)$. Our main structural results follows (for the proof, see [Appendix E.1](#)).

364 **Proposition 1** (Quasar Convexity). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that [Assumption 1](#)
 365 holds. The function \mathcal{L}_λ of Equation (1) with domain \mathcal{W} is $\text{poly}(\epsilon, \alpha, 1/C, 1/D_S, 1/D_{\mathcal{I}})$ -quasar
 366 convex with respect to $-M/\lambda$ on the domain \mathcal{W} .*

367 Since ρ^* is small (by [Proposition 2](#)), $H(\rho^* W)$ is essentially constant and close in value to the negative
 368 entropy of the uniform distribution. Hence, the effect of $R(W)$ during optimization is essentially
 369 the same as that of $H(W)$ (since β^* is close to 0). We show that \mathcal{L}_λ is quasr convex with a non-trivial
 370 parameter γ (see [Proposition 1](#)). We can then apply (in a black-box manner) the convergence results
 371 from [\[HMR16\]](#) to optimize it using projected SGD. We show that SGD finds a weight matrix \widehat{W} such
 372 that the solution generator $p(\widehat{W})$ generates solutions achieving actual loss \mathcal{L} close to that of the near
 373 optimal matrix $\bar{W} = -M/\lambda$, i.e., $\mathcal{L}(\widehat{W}) \leq \mathcal{L}(\bar{W}) + \epsilon$. For further details, see [Appendix E.3](#).

374 5 Experimental Evaluation

375 In this section, we investigate experimentally the effect of our main theoretical contributions, the
 376 entropy regularizer (see Equation (2)) and the fast/slow mixture scheme (see Equation (5)). We try to
 377 find the Max-Cut of a fixed graph G , i.e., the support of the prior \mathcal{R} is a single graph. Similarly to our
 378 theoretical results, our sampler is of the form $e^{\text{score}(s; w)}$, where $s \in \{-1, 1\}^n$ (here n is the number
 379 of nodes in the graph) is a candidate solution of the Max-Cut problem. For the score function we
 380 used a simple linear layer (left plot of [Figure 2](#)) and a 3-layer ReLU network (right plot of [Figure 2](#)).

381 Focusing on instances where the number of nodes n is small (say $n = 15$), we can explicitly compute
 382 the density function and work with an *exact* sampler. We generate 100 random $G(n, p)$ graphs with
 383 $n = 15$ nodes and $p = 0.5$ and train solution generators using both the "vanilla" loss \mathcal{L} and the
 384 entropy-regularized loss \mathcal{L}_λ with the fast/slow mixture scheme. We perform 600 iterations and, for
 385 the entropy regularization, we progressively decrease the regularization weight, starting from 10, and
 386 dividing it by 2 every 60 iterations. Out of the 100 trials we found that our proposed objective was
 387 always able to find the optimal cut while the model trained with the vanilla loss was able to find it for
 388 approximately 65% of the graphs (for 65 out of 100 using the linear network and for 66 using the
 389 ReLU network).

390 Hence, our experiments demonstrate that while the unregularized objective is often "stuck" at sub-
 391 optimal solutions – and this happens even for very small instances ($n = 15$ nodes) – of the Max-Cut
 392 problem, the objective motivated by our theoretical results is able to find the optimal solutions. We
 393 leave more extensive experimental evaluation as an interesting direction for future work. For further
 394 details, see [Appendix I](#).

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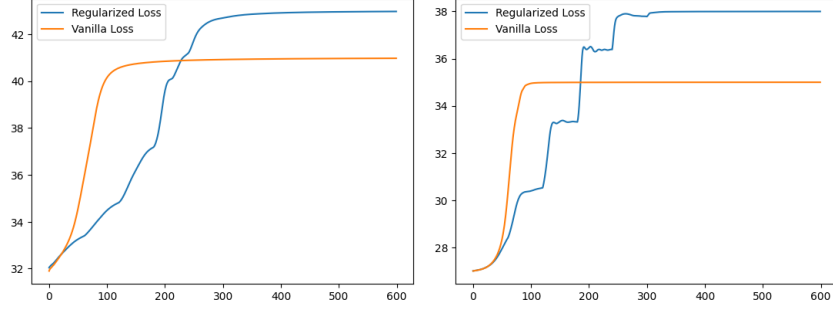


Figure 2: Plot of the Max-Cut value trajectory of the “vanilla” objective and entropy-regularized objective with the slow/fast mixture scheme. We remark that we plot the value of the cut of each iteration (and not the value of the regularized-loss). On the horizontal axis we plot the number of iterations and on the vertical axis we plot the achieved value of the cut. Both graphs used were random $G(n, p)$ graphs generated with $n = 15$ nodes and edge probability $p = 0.5$. For the left plot we used a linear network (the same exponential family as the one used in our theoretical results). For the right plot we used a simple 3-Layer ReLU network to generate the scores. We observe that the “vanilla” loss gets stuck on sub-optimal solutions.

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615 A Preliminaries and Notation

616 This lemma is a useful tool for quasar convex functions.

617 **Lemma 1** ([HMR16]). *Suppose that the functions f_1, \dots, f_n are individually γ -quasar convex in X*
 618 *with respect to a common global minimum \bar{x} . Then for non-negative weights a_1, \dots, a_n , the linear*
 619 *combination $f = \sum_{i \in [n]} a_i f_i$ is also γ -quasar convex with respect to \bar{x} in X .*

620 In the proofs, we use the following notation: for a matrix W and vectors x, z , we let

$$\phi(x; z; W) = \frac{\exp(z^\top W z)}{\sum_{y \in X} \exp(z^\top W y)}, \quad (1)$$

621 be a probability mass function over X and we overload the notation as

$$\phi(x; w) = \frac{\exp(w \cdot x)}{\sum_{y \in X} \exp(w \cdot y)}. \quad (2)$$

622 B The Proof of Remark 1

623 *Proof.* Let x_1, \dots, x_n be the variables of the Max-Cut problem of interest and $S = \{-1, 1\}^n$ be
 624 the solution space. Consider \mathcal{P} to be the collection of product distributions over S , i.e., for any
 625 $p \in \mathcal{P}$, it holds that, for any $s \in S$, $\Pr_{x \sim p}[x = s] = \prod_{i \in [n]} p_i^{\frac{1+s_i}{2}} (1-p_i)^{\frac{1-s_i}{2}}$. Let us consider
 626 the cube $[\epsilon, 1-\epsilon]^n$. This family is complete since the $O(\epsilon)$ -sub-optimal solution of I belongs to \mathcal{P}
 627 and is compressed since the description size is $\text{poly}(n, \log(1/\epsilon))$. We show that in this setting there
 628 exist bad stationary points. Let L_G be the Laplacian matrix of the input graph. For some product
 629 distribution $p \in \mathcal{P}$, it holds that

$$\mathcal{L}(p) = -\mathbf{E}_{x \sim p(\cdot)}[x^\top L_G x] = -(2p-1)^\top \bar{L}_G (2p-1), \quad \nabla_p \mathcal{L}(p) = -4\bar{L}_G (2p-1),$$

630 where \bar{L}_G is zero in the diagonal and equal to the Laplacian otherwise. Let us consider a vertex of the
 631 cube $p \in [\epsilon, 1-\epsilon]^n$ which is highly and strictly sub-optimal, i.e., any single change of a node would
 632 strictly improve the number of edges in the cut and the score attained in p is very large compared to
 633 $\min_{x \in S} -x^\top L_G x$. For any $i \in [n]$, we show that

$$(\nabla \mathcal{L}(p) \cdot e_i)((2p-1) \cdot e_i) < 0.$$

634 This means that if p_i is large (i.e., $1-\epsilon$), then the i -th coordinate of the gradient of $\mathcal{L}(p)$ should be
 635 negative since this would imply that the negative gradient would preserve p_i to the right boundary.
 636 Similarly for the case where p_i is small. This means that this point is a stationary point and is highly
 637 sub-optimal by assumption.

638 Let P (resp. N) be the set of indices in $[n]$ where p takes the value $1-\epsilon$ (resp. ϵ). For any $i \in [n]$,
 639 let $\mathcal{N}(i)$ be its neighborhood in G . Let us consider $i \in P$. We have that $(2p-1) \cdot e_i > 0$ and so it
 640 suffices to show that

$$(\bar{L}_G (2p-1)) \cdot e_i > 0,$$

641 which corresponds to showing that

$$\sum_{j \in \mathcal{N}(i) \cap P} L_G(i, j)(1-2\epsilon) + \sum_{j \in \mathcal{N}(i) \cap N} L_G(i, j)(2\epsilon-1) > 0,$$

642 and so we would like to have

$$\sum_{j \in P} L_G(i, j) - \sum_{j \in N} L_G(i, j) > 0.$$

643 Note that this is true for any $i \in [n]$ since the current solution is a strict local optimum. The same
 644 holds if $i \in N$. \square

C Completeness

Proposition 2 (Completeness). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that [Assumption 1](#) holds. There exist $\beta^*, \rho^* \in (0, 1)$ and \mathcal{W} such that the family of solution generators \mathcal{P} of [Equation \(5\)](#) is complete.*

Proof. Assume that $\mathcal{O}(s, I) = \psi_{\mathcal{I}}(I)^\top M \psi_S(s)$ and let $z = \psi_{\mathcal{I}}(I)$ and $x = \psi_S(s)$. Moreover, let $\alpha, C, D_S, D_{\mathcal{I}}$ be the parameters promised by [Assumption 1](#). Let us consider the family $\mathcal{P} = \{p(W) : W \in \mathcal{W}\}$ with

$$p(x; z; W) = (1 - \beta^*) \frac{e^{z^\top W x}}{\sum_{y \in X} e^{z^\top W y}} + \beta^* \frac{e^{z^\top \rho^* W x}}{\sum_{y \in X} e^{z^\top \rho^* W y}},$$

where the mixing weight $\beta^* \in (0, 1)$ and the inverse temperate ρ^* are to be decided. Recall that

$$\mathcal{L}(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; W)} [L(x; z)] = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; W)} [z^\top M x].$$

Let us pick the parameter matrix $W = -M/\lambda$. Let us now fix a $z \in \psi_{\mathcal{I}}(\mathcal{I})$. For the given matrix M , we can consider the finite set of values V obtained by the quadratic forms $\{z^\top M x\}_{x \in \psi_S(S)}$. We further cluster these values so that they have distance at least ϵ between each other. We consider the level sets C_i where C_1 is the subset of S with minimum value $v_1 (= v_1(z)) \in V$, C_2 is the subset with the second smallest $v_2 (= v_2(z)) \in V$, etc. For fixed $z \in \psi_{\mathcal{I}}(\mathcal{I})$, we have that

$$\mathbf{E}_{x \sim p(\cdot; z; -M/\lambda)} [z^\top M x] = (1 - \beta^*) \mathbf{E}_{x \sim \phi(\cdot; z; -M/\lambda)} [z^\top M x] + \beta^* \mathbf{E}_{x \sim \phi(\cdot; z; -\rho^* M/\lambda)} [z^\top M x],$$

where ϕ comes from [\(1\)](#). We note that

$$\Pr_{x \sim \phi(\cdot; z; -M/\lambda)} [z^\top M x \in C_i] = \frac{|C_i| e^{-v_i/\lambda}}{\sum_j |C_j| e^{-v_j/\lambda}}.$$

We claim that, by letting $\lambda \rightarrow 0$, the above measure concentrates uniformly on C_1 . The worst case scenario is when $|C_2| = |S| - |C_1|$ and $v_2 = v_1 + \epsilon$. Then we have that

$$\Pr_{x \sim \phi(\cdot; z; -M/\lambda)} [z^\top M x \in C_2] = \frac{|C_2|/|C_1| e^{(-v_2+v_1)/\lambda}}{1 + |C_2|/|C_1| e^{(-v_2+v_1)/\lambda}} \leq \delta,$$

when $1/\lambda > \log(|\psi_S(S)|/\delta)/\epsilon$, since in the worst case $|C_2|/|C_1| = \Omega(|\psi_S(S)|)$. Using this choice of λ and taking expectation over z , we get that

$$\mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; -M/\lambda)} [z^\top M x] \leq (1 - \beta^*) \mathbf{E}_{z \sim \mathcal{R}} \left[(1 - \delta) \min_{x \in \psi_S(S)} L(x; z) + \delta v_2(z) \right] + \beta^* \mathbf{E}_{x \sim \phi(\cdot; z; -\rho^* M/\lambda)} [z^\top M x].$$

First, we remark that by taking $\rho^* = \text{poly}(\alpha, 1/C, 1/D_S, 1/D_{\mathcal{I}})$, the last term in the right-hand side of the above expression can be replaced by the expected score of an almost-uniform solution (see [Lemma 3](#) and [Proposition 8](#)), which is at most $\text{poly}(D_S, D_{\mathcal{I}}, C) 2^{-|\psi_S(S)|}$ (and which is essentially negligible). Finally, one can pick $\beta^*, \delta = \text{poly}(\epsilon, 1/C, 1/D_S, 1/D_{\mathcal{I}})$ so that

$$\mathcal{L}(-M/\lambda) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; -M/\lambda)} [z^\top M x] \leq \mathbf{E}_{z \sim \mathcal{R}} \left[\min_{x \in \psi_S(S)} L(x; z) \right] + \epsilon.$$

This implies that \mathcal{P} is complete by letting $\bar{W} = -M/\lambda \in \mathcal{W}$. This means that one can take \mathcal{W} be a ball centered at 0 with radius (of ϵ -sub-optimality) to be of order at least $B = \|M\|_{\text{F}}/\lambda$. \square

D Compression

Proposition 3 (Compression). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that [Assumption 1](#) holds. There exist $\beta^*, \rho^* \in (0, 1)$ and \mathcal{W} such that the family of solution generators \mathcal{P} of [Equation \(5\)](#) is compressed.*

Proof. We have that the bit complexity to represent the mixing weight β^* is $\text{polylog}(D_S, D_{\mathcal{I}}, C, 1/\epsilon)$ and the description size of \mathcal{W} is polynomial in $|\mathcal{I}|$ and in $\log(1/\epsilon)$. This follows from [Assumption 1](#) since the feature dimensions n_X and n_Z are $\text{poly}(|\mathcal{I}|)$ and \mathcal{W} is a ball centered at 0 with radius $O(B)$, where $B = \|M\|_{\text{F}}/\lambda \leq C/\lambda$, which are also $\text{poly}(|\mathcal{I}|/\epsilon)$. \square

677 E Efficiently Optimizable

678 **Proposition 4** (Efficiently Optimizable). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that*
 679 *Assumption 1 holds. There exist $\beta^*, \rho^* \in (0, 1)$ and \mathcal{W} such that family of solution generators \mathcal{P} of*
 680 *Equation (5) is efficiently optimizable using Projected SGD, where the projection set is \mathcal{W} .*

681 The proof of this proposition is essentially decomposed into two parts: first, we show that the
 682 entropy-regularized loss of Equation (2) is quasar convex and then apply the projected SGD algorithm
 683 to \mathcal{L}_λ .

684 Recall that $H(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim \phi(\cdot; z; W)} [\log \phi(x; z; W)]$. Let R be a weighted sum (to be in
 685 accordance with the mixture structure of \mathcal{P}) of negative entropy regularizers

$$R(W) = (1 - \beta^*)H(W) + \frac{\beta^*}{\rho^*}H(\rho^*W), \quad (1)$$

686 where β^*, ρ^* are the fixed parameters of \mathcal{P} (recall Equation (5)). We define the regularized loss

$$\mathcal{L}_\lambda(W) = \mathcal{L}(W) + \lambda R(W), \quad (2)$$

687 where

$$\mathcal{L}(W) = \mathbf{E}_{z \sim \mathcal{R}} \mathbf{E}_{x \sim p(\cdot; z; W)} [L(z; x)], \quad p(W) \in \mathcal{P}.$$

688 E.1 Quasar Convexity of the Regularized Loss

689 In this section, we show that \mathcal{L}_λ of Equation (2) is quasar convex. We restate Proposition 1.

690 **Proposition 5** (Quasar Convexity). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that Assumption 1*
 691 *holds. The function \mathcal{L}_λ of Equation (2) with domain \mathcal{W} is $\text{poly}(C, D_S, D_{\mathcal{I}}, 1/\epsilon, 1/\alpha)$ -quasar convex*
 692 *with respect to $-M/\lambda$ on the domain \mathcal{W} .*

693 *Proof.* We can write the loss \mathcal{L}_λ as

$$\mathcal{L}_\lambda(W) = \mathbf{E}_{z \sim \mathcal{R}} [\mathcal{L}_{\lambda, z}(W)] = \mathbf{E}_{z \sim \mathcal{R}} [\mathcal{L}_z(W) + \lambda R_z(W)],$$

694 where the mappings \mathcal{L}_z and R_z are instance-specific (i.e., we have fixed z). We can make use of
 695 Lemma 1, which states that linear combinations of quasar convex (with the same minimizer) remain
 696 quasar convex. Hence, since the functions $\mathcal{L}_{\lambda, z}$ have the same minimizer $-M/\lambda$, it suffices to show
 697 quasar convexity for a particular fixed instance mapping, i.e., it suffices to show that the function

$$\mathcal{L}_{\lambda, z}(W) = \mathcal{L}_z(W) + \lambda R_z(W)$$

698 is quasar convex. Recall that W is a matrix of dimension $n_Z \times n_X$. To deal with the function $\mathcal{L}_{\lambda, z}$,
 699 we consider the simpler function that maps vectors instead of matrices to real numbers. For some
 700 vector c , let $\mathcal{L}_\lambda^{\text{vec}} : \mathbb{R}^{n_X} \rightarrow \mathbb{R}$ be

$$\mathcal{L}_\lambda^{\text{vec}}(w) = \mathbf{E}_{x \sim p(\cdot; w)} [c \cdot x] + \lambda R^{\text{vec}}(w), \quad (3)$$

701 where for any vector $w \in \mathbb{R}^{n_X}$, we define the probability distribution $\phi(\cdot; w)$ over the solution space
 702 $X = \psi_S(S)$ with probability mass function

$$\phi(x; w) = \frac{e^{w \cdot x}}{\sum_{y \in X} e^{w \cdot y}}.$$

703 We then define

$$p(\cdot; w) = (1 - \beta^*)\phi(\cdot; w) + \beta^*\phi(\cdot; \rho^*w),$$

704 and $R^{\text{vec}}(w) = (1 - \beta^*)H(w) + \frac{\beta^*}{\rho^*}H(\rho^*w)$ (this is essentially a weighted sum of regularizers,
 705 needed to simplify the proof) with $H(w) = \mathbf{E}_{x \sim \phi(\cdot; w)} \log \phi(x, w)$. These quantities are essentially
 706 the fixed-instance analogues of Equations (5) and (2). The crucial observation is that by taking
 707 $c = z^\top M$ and applying the chain rule we have that

$$\nabla_W \mathcal{L}_{\lambda, z}(W) = z \cdot [\nabla_w \mathcal{L}_\lambda^{\text{vec}}(z^\top W)]^\top. \quad (4)$$

This means that the gradient of the fixed-instance objective $\mathcal{L}_{\lambda,z}$ is a matrix of dimension $n_Z \times n_X$ that is equal to the outer product of the instance featurization z and the gradient of the simpler function $\mathcal{L}_w^{\text{vec}}$ evaluated at $z^\top W$. Let us now return on showing that $\mathcal{L}_{\lambda,z}$ is quasar convex. To this end, we observe that

$$\nabla_W \mathcal{L}_{\lambda,z}(W) \cdot \left(W + \frac{M}{\lambda}\right) = \nabla_w \mathcal{L}_\lambda^{\text{vec}}(z^\top W) \cdot \left(z^\top W + z^\top \frac{M}{\lambda}\right).$$

This means that, since z is fixed, it suffices to show that the function $\mathcal{L}_\lambda^{\text{vec}}$ is quasar convex. We provide the next key proposition that deals with issue. This result is one the main technical aspects of this work and its proof can be found in [Appendix E.2](#).

In the following, intuitively \mathcal{X} is the post-featurization instance space and \mathcal{Z} is the parameter space.

Proposition 6. Consider $\epsilon, \lambda > 0$. Let $\|c\|_2 \leq C_1$. Let \mathcal{Z} be an open ball centered at 0 with diameter $2C_1/\lambda$. Let \mathcal{X} be a space of diameter D and let $\mathbf{Var}_{x \sim U(\mathcal{X})}[v \cdot x] \geq \alpha \|v\|_2^2$ for any $v \in \mathcal{Z}$. The function $\mathcal{L}_\lambda^{\text{vec}}(w) = \mathbf{E}_{x \sim p(\cdot; w)}[c \cdot x] + \lambda R^{\text{vec}}(w)$ is $\text{poly}(1/C_1, 1/D, \epsilon, \alpha)$ -quasar convex with respect to $-c/\lambda$ on \mathcal{Z} .

We can apply the above result with $c = z^\top M$, $w = z^\top W$, $D = D_S$ and $C_1 = D_{\mathcal{I}}C$. These give that the quasar convexity parameter γ is of order $\gamma = \text{poly}(\epsilon, \alpha, 1/C, 1/D_S, 1/D_{\mathcal{I}})$. Since we have that $\mathcal{L}_\lambda^{\text{vec}}(z^\top W) = \mathcal{L}_{\lambda,z}(W)$, we get that

$$\nabla_W \mathcal{L}_{\lambda,z}(W) \cdot (W + M/\lambda) \geq \gamma(\mathcal{L}_{\lambda,z}(W) - \mathcal{L}_{\lambda,z}(-M/\lambda)).$$

This implies that $\mathcal{L}_{\lambda,z}$ is γ -quasar convex with respect to the minimizer $-M/\lambda$ and completes the proof using [Lemma 1](#). \square

E.2 The Proof of Proposition 6

Let us consider $\mathcal{L}_\lambda^{\text{vec}}$ to be a real-valued differentiable function defined on \mathcal{Z} . Let $w, -c/\lambda \in \mathcal{Z}$ and let L be the line segment between them with $L \in \mathcal{Z}$. The mean value theorem implies that there exists $w' \in L$ such that

$$\mathcal{L}_\lambda^{\text{vec}}(w) - \mathcal{L}_\lambda^{\text{vec}}(-c/\lambda) = \nabla_w \mathcal{L}_\lambda^{\text{vec}}(w') \cdot (w + c/\lambda) \leq \|\nabla \mathcal{L}_\lambda^{\text{vec}}(w')\|_2 \|w + c/\lambda\|_2.$$

Now we have that $\mathcal{L}_\lambda^{\text{vec}}$ has bounded gradient (see [Lemma 2](#)) and so we get that

$$\|\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w')\|_2 \leq D^2 \|c + \lambda w'\|_2 = D^2 \lambda \|w' + c/\lambda\|_2 \leq D^2 \lambda \|w + c/\lambda\|_2,$$

since $w' \in L$. This implies that

$$\mathcal{L}_\lambda^{\text{vec}}(w) - \mathcal{L}_\lambda^{\text{vec}}(-c/\lambda) \leq D^2 \lambda \|w + c/\lambda\|_2^2 \leq \frac{1}{\gamma} \nabla \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (w + c/\lambda),$$

where $1/\gamma = \frac{\text{poly}(C_1, D)}{\epsilon^3 \alpha^2}$. The last inequality is an application of the correlation lower bound (see [Lemma 3](#)).

In the above proof, we used two key lemmas: a bound for the norm of the gradient and a lower bound for the correlation. In the upcoming subsections, we prove these two results.

E.2.1 Bounded Gradient Lemma and Proof

Lemma 2 (Bounded Gradient Norm of $\mathcal{L}_\lambda^{\text{vec}}$). Consider $\epsilon, \lambda > 0$. Let \mathcal{Z} be the domain of $\mathcal{L}_\lambda^{\text{vec}}$ of (3). Let \mathcal{X} be a space of diameter D . For any $w \in \mathcal{Z}$, it holds that

$$\|\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w)\|_2 \leq O(D^2) \|c + \lambda w\|_2.$$

Proof. We have that

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) = (1 - \beta^*) G_w + \beta^* \rho^* G_{\rho^* w},$$

where

$$G_w = \mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; w)}[x],$$

740 and

$$G_{\rho^* w} = \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[x].$$

741 Note that since $x \in \mathcal{X}$, it holds that $\|x\|_2 \leq D$. Hence

$$\|G_w\|_2 = \sup_{v: \|v\|_2=1} |v \cdot G_w| \leq 2D^2 \|c + \lambda w\|_2.$$

742 Moreover, we have that

$$\|G_{\rho^* w}\|_2 \leq 2D^2 \|c + \lambda w\|_2.$$

743 This means that

$$\|\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w)\|_2 \leq 2(1 - \beta^*) \|c + \lambda w\|_2 D^2 + 2\beta^* \rho^* \|c + \lambda w\|_2 D^2 = O(D^2) \|c + \lambda w\|_2.$$

744

□

745 E.2.2 Correlation Lower Bound Lemma and Proof

746 The following lemma is the second ingredient in order to show [Proposition 6](#).

747 **Lemma 3** (Correlation Lower Bound for $\mathcal{L}_\lambda^{\text{vec}}$). *Let $\lambda > 0$. Let $\|c\|_2 \leq C_1$. Let \mathcal{Z} be an open ball*
 748 *centered at 0 with diameter $B = 2C_1/\lambda$. Let \mathcal{X} be a space of diameter D . Assume that $w \in \mathcal{Z}$ and*
 749 *$\mathbf{Var}_{x \sim U(\mathcal{X})}[(c + \lambda w) \cdot x] \geq \alpha \|c + \lambda w\|_2^2$ for some $\alpha > 0$. Then, for any $\beta^* \in (0, 1)$, there exists*
 750 *$\rho^* > 0$ such that it holds that*

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) = \Omega(\beta^* \alpha^2 / (BD^3) \|c + \lambda w\|_2^2),$$

751 *where $\mathcal{L}_\lambda^{\text{vec}}$ is the regularized loss of [Proposition 6](#), ρ^* is the scale in the second component of the*
 752 *mixture of (5) and $\beta^* \in (0, 1)$ is the mixture weight.*

753 First, in [Lemma 4](#) and [Lemma 5](#), we give a formula for the desired correlation $\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w)$
 754 and, then we can provide a proof for [Lemma 3](#) by lower bounding this formula.

755 **Lemma 4** (Correlation with Regularization). *Consider the function $g(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[c \cdot x] + \lambda H(w)$,*
 756 *where H is the negative entropy regularizer. Then it holds that*

$$\nabla_w g(w) \cdot (c + \lambda w) = \mathbf{Var}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x].$$

757 *Proof.* Let us consider the following objective function:

$$g(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[c \cdot x] + \lambda H(w), \quad \phi(x; w) = \frac{\exp(w \cdot x)}{\sum_{y \in \mathcal{X}} \exp(w \cdot y)},$$

758 where H is the negative entropy regularizer, i.e.,

$$H(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[\log \phi(x; w)] = \mathbf{E}_{x \sim \phi(\cdot; w)}[w \cdot x] - \log \left(\sum_{y \in \mathcal{X}} e^{w \cdot y} \right).$$

759 The gradient of g with respect to $w \in \mathcal{W}$ is equal to

$$\nabla_w g(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[(c \cdot x) \nabla_w \log \phi(x; w)] + \lambda \nabla_w H(w).$$

760 It holds that

$$\nabla_w \log \phi(x; w) = \nabla_w \left(w \cdot x - \log \sum_{y \in \mathcal{X}} e^{w \cdot y} \right) = x - \mathbf{E}_{x \sim \phi(\cdot; w)}[x],$$

761 and

$$\nabla H(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[x] + \mathbf{E}_{x \sim \phi(\cdot; w)}[(w \cdot x) \nabla_w \log \phi(x; w)] - \mathbf{E}_{x \sim \phi(\cdot; w)}[x] = \mathbf{E}_{x \sim \phi(\cdot; w)}[(w \cdot x) \nabla_w \log \phi(x; w)].$$

762 So, we get that

$$\nabla_w g(w) = \mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; w)}[x].$$

763 Note that

$$\nabla_w g(w) \cdot (c + \lambda w) = \mathbf{Var}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x].$$

764

□

Lemma 5 (Gradient with Regularization and Mixing). *For any $\epsilon > 0$, for the family of solution generators $\mathcal{P} = \{p(\cdot; w) = (1 - \beta^*)\phi(\cdot; w) + \beta^*\phi(\cdot; \rho^*w) : w \in \mathcal{W}\}$ and the objective $\mathcal{L}_\lambda^{\text{vec}}$ of Equation (3), it holds that*

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) = (1 - \beta^*) \mathbf{Var}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] + \beta^* \rho^* \mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x],$$

for any $\lambda > 0$.

Proof. Let us first consider the scaled parameter $\rho w \in \mathcal{W}$ for some $\rho > 0$. Then it holds that

$$\nabla_w \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[c \cdot x] = \mathbf{E}_{x \sim \phi(\cdot; \rho w)} \left[(c \cdot x) \left(\rho x - \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[\rho x] \right) \right] = \rho \mathbf{E}_{x \sim \phi(\cdot; \rho w)} \left[(c \cdot x) \left(x - \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[x] \right) \right].$$

Moreover, the negative entropy regularizer at ρw is

$$H(\rho w) = \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(\rho w) \cdot x] - \log \sum_{y \in \mathcal{X}} e^{(\rho w) \cdot y}.$$

It holds that

$$\nabla_w H(\rho w) = \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(\rho w \cdot x) \nabla_w \log \phi(x; \rho w)] = \rho^2 \left(\mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(w \cdot x)x] - \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[w \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[x] \right).$$

We consider the objective function $\mathcal{L}_\lambda^{\text{vec}}$ to be defined as follows: first, we take

$$p(\cdot; w) = (1 - \beta^*)\phi(\cdot; w) + \beta^*\phi(\cdot; \rho^*w),$$

i.e., $p(\cdot; w)$ is the mixture of the probability measures $\phi(\cdot; w)$ and $\phi(\cdot; \rho^*w)$ with weights $1 - \beta^*$ and

β^* respectively for some scale $\rho^* > 0$. Moreover, we take $R^{\text{vec}}(w) = (1 - \beta^*)H(w) + \frac{\beta^*}{\rho^*}H(\rho^*w)$.

Then we define our regularized loss $\mathcal{L}_\lambda^{\text{vec}}$ to be

$$\mathcal{L}_\lambda^{\text{vec}}(w) = \mathbf{E}_{x \sim p(\cdot; w)}[c \cdot x] + \lambda R^{\text{vec}}(w).$$

Using Lemma 4 and the above calculations, we have that

$$\begin{aligned} \nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) &= (1 - \beta^*) \nabla_w \mathbf{E}_{x \sim \phi(\cdot; w)}[c \cdot x] + \lambda(1 - \beta^*) \nabla_w H(w) + \beta^* \nabla_w \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[c \cdot x] + \lambda \frac{\beta^*}{\rho^*} \nabla_w H(\rho^* w) \\ &= (1 - \beta^*) \left(\mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; w)}[x] \right) + \\ &\quad + \beta^* \rho^* \left(\mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[x] \right). \end{aligned}$$

The above calculations yield

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) = (1 - \beta^*) \mathbf{Var}_{x \sim \phi(\cdot; w)}[(c + \lambda w) \cdot x] + \beta^* \rho^* \mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x],$$

and this concludes the proof. \square

The above correlation being positive intuitively means that performing gradient descent to $\mathcal{L}_\lambda^{\text{vec}}$ gives that the parameter w converges to $-c/\lambda$, the point that achieves completeness for that objective.

However, to obtain fast convergence, we need to show that the above correlation is non-trivial. This means that our goal in order to prove Lemma 3 is to provide a lower bound for the above quantity, i.e., it suffices to give a non-trivial lower bound for the variance of the random variable $(c + \lambda w) \cdot x$ with respect to the probability measure $\phi(\cdot; \rho^*w)$. It is important to note that in the above statement we did not fix the value of ρ^* . We can now make use of Proposition 8. Intuitively, by taking the scale parameter appearing in the mixture ρ^* to be sufficiently small, we can manage to provide a lower bound for the variance of $(c + \lambda w) \cdot x$ with respect to the almost uniform measure, i.e., the second summand of the above right-hand side expression has significant contribution. We remark that ρ corresponds to the inverse temperature parameter. Hence, our previous analysis essentially implies that policy gradient on combinatorial optimization potentially works if the variance $\mathbf{Var}_{x \sim \phi(\cdot; \rho w)}[(c + \lambda w) \cdot x]$ is non-vanishing at high temperatures $1/\rho$.

792 The proof of [Lemma 3](#). Recall that

$$\phi(x; w) = \frac{e^{w \cdot x}}{\sum_{y \in \mathcal{X}} e^{w \cdot y}}.$$

793 Let also D be the diameter of \mathcal{X} and B the diameter of \mathcal{Z} . Recall from [Lemma 5](#) that we have that

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) = (1 - \beta^*) \mathbf{Var}_{x \sim p(\cdot; w)}[(c + \lambda w) \cdot x] + \beta^* \rho^* \mathbf{Var}_{x \sim p(\cdot; \rho^* w)}[(c + \lambda w) \cdot x],$$

794 where the scale parameter $\rho^* > 0$ is to be decided. This means that

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) \geq \beta^* \rho^* \mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x].$$

795 Our goal is now to apply [Proposition 8](#) in order to lower bound the above variance. Applying
796 [Proposition 8](#) for $\mu \leftarrow \phi(\cdot; \rho^* w)$, $c \leftarrow c + \lambda w \in \mathcal{Z}$ and, so for some absolute constant C_0 , we can
797 pick

$$\rho^* = C_0 \frac{\alpha}{BD^3}.$$

798 Thus, we have that

$$\mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[(c + \lambda w) \cdot x] \geq \Omega(\alpha \|c + \lambda w\|_2^2).$$

799 This implies the desired result since

$$\nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) \cdot (c + \lambda w) \geq C_0 \beta^* \alpha^2 / (BD^3) \|c + \lambda w\|_2^2.$$

800 □

801 E.3 Convergence for Quasar Convex Functions

802 The fact that \mathcal{L}_λ is quasar convex with respect to $-M/\lambda$ implies that projected SGD converges to
803 that point in a small number of steps and hence the family \mathcal{P} is efficiently optimizable. The analysis
804 is standard (see e.g., [\[HMR16\]](#)). For completeness a proof can be found in [Appendix E.3](#).

805 **Proposition 7** (Convergence). *Consider $\epsilon > 0$ and a prior \mathcal{R} over \mathcal{I} . Assume that [Assumption 1](#)
806 holds with parameters $C, D_S, D_{\mathcal{I}}, \alpha$. Let W_1, \dots, W_T be the updates of the SGD algorithm with
807 projection set \mathcal{W} performed on \mathcal{L}_λ of Equation (2) with appropriate step size and parameter λ . Then,
808 for the non-regularized objective \mathcal{L} , it holds that*

$$\mathbf{E}_{t \sim U([T])} [\mathcal{L}(W_t)] \leq \mathcal{L}(-M/\lambda) + \epsilon,$$

809 when $T \geq \text{poly}(1/\epsilon, 1/\alpha, C, D_S, D_{\mathcal{I}}, \|W_0 + M/\lambda\|_{\mathcal{F}})$.

810 Our next goal is to use [Proposition 1](#) and show that standard projected SGD on the objective \mathcal{L}_λ
811 converges in a polynomial number of steps. The intuition behind this result is that since the correlation
812 between $\nabla \mathcal{L}_\lambda(W)$ and the direction $M + \lambda W$ is positive and non-trivial, the gradient field drives the
813 optimization method towards the point $-M/\lambda$.

814 *Proof.* Consider the sequence of matrices $W_1, \dots, W_t, \dots, W_T$ generated by applying PSGD on \mathcal{L}_λ
815 with step size η (to be decided) and initial parameter vector $W_0 \in \mathcal{W}$. We have that \mathcal{L}_λ is γ -quasar
816 convex and is also $O(\Gamma)$ -weakly smooth² since we now show that it is Γ -smooth.

817 **Lemma 6.** \mathcal{L}_λ is $\text{poly}(D_S, D_{\mathcal{I}}, C)$ -smooth.

818 *Proof.* We have that

$$\|\nabla \mathcal{L}_\lambda(W)\|_{\mathcal{F}}^2 = \left\| \mathbf{E}_{z \sim \mathcal{R}} [\nabla \mathcal{L}_{\lambda, z}(W)] \right\|_{\mathcal{F}}^2 \leq \mathbf{E}_{z \sim \mathcal{R}} \|z\|_2^2 \|\nabla_w \mathcal{L}_\lambda^{\text{vec}}(z^\top W)\|_{\mathcal{F}}^2.$$

819 It suffices to show that $\mathcal{L}_\lambda^{\text{vec}}$ is smooth. Recall that

$$\begin{aligned} \nabla_w \mathcal{L}_\lambda^{\text{vec}}(w) &= (1 - \beta^*) \left(\mathbf{E}_{x \sim \phi(\cdot; w)} [(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; w)} [(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; w)} [x] \right) + \\ &\quad + \beta^* \rho^* \left(\mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [(c + \lambda w) \cdot x] - \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [x] \right). \end{aligned}$$

²As mentioned in [\[HMR16\]](#), a function f is Γ -weakly smooth if for any point θ , $\|\nabla f(\theta)\|^2 \leq \Gamma(f(\theta) - f(\theta^*))$. Moreover, a function f that is Γ -smooth (in the sense $\|\nabla^2 f\| \leq \Gamma$), is also $O(\Gamma)$ -weakly smooth.

820 This means that

$$\|\nabla_w^2 \mathcal{L}_\lambda^{\text{vec}}(w)\|_F^2 \leq (1 - \beta^*)(A_1 + A_2) + \beta^* \rho^*(A_3 + A_4),$$

821 where

$$\begin{aligned} A_1 &= \left\| \nabla_w \mathbf{E}_{x \sim \phi(\cdot; w)} [(c + \lambda w) \cdot x] \right\|_F^2, \quad A_2 = \left\| \nabla_w \mathbf{E}_{x \sim \phi(\cdot; w)} [(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; w)} [x] \right\|_F^2, \\ A_3 &= \left\| \nabla_w \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [(c + \lambda w) \cdot x] \right\|_F^2, \quad A_4 = \left\| \nabla_w \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [(c + \lambda w) \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)} [x] \right\|_F^2. \end{aligned}$$

823 Standard computation of these values yields that, since D_S and $D_{\mathcal{I}}$ are bounds to x and z respectively,
824 we have that \mathcal{L}_λ is smooth with parameter $\text{poly}(D_S, D_{\mathcal{I}}, C)$. \square

825 Let V be the variance of the unbiased estimator used for $\nabla_W \mathcal{L}_\lambda(W)$. We can apply the next result of
826 [HMR16].

827 **Lemma 7** ([HMR16]). *Suppose the objective function f is γ -weakly quasi convex and Γ -weakly*
828 *smooth, and let $r(\cdot)$ be an unbiased estimator for $\nabla f(\theta)$ with variance V . Moreover, suppose the*
829 *global minimum $\bar{\theta}$ belongs to \mathcal{W} , and the initial point θ_0 satisfies $\|\theta_0 - \bar{\theta}\|_2 \leq R$. Then projected*
830 *stochastic gradient descent with a proper learning rate returns θ_T in T iterations with expected error*

$$\mathbf{E}_{t \sim U([T])} f(\theta_t) - f(\bar{\theta}) \leq \max \left\{ \frac{\Gamma R^2}{\gamma^2 T}, \frac{R\sqrt{V}}{\gamma\sqrt{T}} \right\}.$$

831 We apply the above result to \mathcal{L}_λ in order to find matrices W_1, \dots, W_T that achieve good loss on average
832 compared to $-M/\lambda$. Moreover, using a batch SGD update, we can take V to be also polynomial
833 in the crucial parameters of the problem. We note that one can adapt the above convergence proof
834 and show that the actual loss \mathcal{L} (and not the loss \mathcal{L}_λ) are close after sufficiently many iterations (as
835 indicated by the above lemma). We know that the Frobenius norm of the gradient of $\mathcal{L}(W)$ is at most
836 of order $O(D_{\mathcal{I}}^2 C D_S^2)$. We can apply the mean value theorem in high dimensions (by taking \mathcal{W} to
837 be an open ball of radius $O(B)$) and this yields that the difference between the values of $\mathcal{L}(W_T)$
838 and $\mathcal{L}(-M/\lambda)$ is at most $D_{\mathcal{I}}^2 C D_S^2 \|W_t + M/\lambda\|_F^2$. However, the right-hand side is upper bounded
839 by the correlation between $\nabla \mathcal{L}_\lambda(W_t)$ and $W_t + M/\lambda$. Hence, we can still use this correlation as a
840 potential in order to minimize \mathcal{L} . This implies that the desired convergence guarantee holds as long
841 as $T \geq \text{poly}(1/\epsilon, 1/\alpha, C, D_S, D_{\mathcal{I}}, \|W_0 + M/\lambda\|_F)$. \square

842 F Deferred Proofs: Variance under Almost Uniform Distributions

843 This section is a technical section that states some properties of exponential families. We use some
844 standard notation, such as w and x , for the statements and the proofs but we underline that these
845 symbols do not correspond to the notation in the main body of the paper.

846 We consider the parameter space Θ and for any parameter $w \in \Theta$, we define the probability
847 distribution $\phi(\cdot; w)$ over a space \mathcal{X} with density

$$\phi(x; w) = \frac{e^{w \cdot x}}{\sum_{y \in \mathcal{X}} e^{w \cdot y}}.$$

848 In this section, our goal is to relate the variance of $c \cdot x$ under the measure $\phi(\cdot; 0)$ (uniform case) and
849 $\phi(\cdot; \rho^* w)$ for some $w \in \mathcal{W}$ and some sufficiently small ρ^* (almost uniform case). The main result of
850 this section follows.

851 **Proposition 8** (Variance Lower Bound Under Almost Uniform Distributions). *Assume that the*
852 *variance of $c \cdot x$ under the uniform distribution over \mathcal{X} , whose diameter is D , is lower bounded by*
853 *$\alpha \|c\|_2^2$. Moreover assume that $w \in \Theta$ with $\|w\|_2 \leq B$. Then, setting $\rho^* = O(\alpha/(BD^3))$, it holds*
854 *that $\text{Var}_{x \sim \phi(\cdot; \rho^* w)}[c \cdot x] = \Omega(\alpha \|c\|_2^2)$.*

855 We first provide a general abstract lemma that relates the variance of the uniform distribution U over
856 \mathcal{X} to the variance of an almost uniform probability measure μ . For simplicity, we denote the uniform
857 distribution over \mathcal{X} with $U = U(\mathcal{X})$.

Lemma 8. Let $w \in \Theta$ and $x \in \mathcal{X}$ with $\|x\|_2 \leq D$. Consider the uniform probability measure U over \mathcal{X} and let μ over \mathcal{X} be such that there exist $\epsilon_1, \epsilon_2 > 0$ with:

- $\|\mathbf{E}_{x \sim U}[x] - \mathbf{E}_{x \sim \mu}[x]\|_2 \leq \epsilon_1$, and,
- $w^\top \mathbf{E}_{x \sim \mu}[xx^\top]w \geq w^\top \mathbf{E}_{x \sim U}[xx^\top]w - \epsilon_2 \|w\|_2^2$.

Then it holds that $\text{Var}_{x \sim \mu}[w \cdot x] \geq \text{Var}_{x \sim U}[w \cdot x] - 3 \max\{\epsilon_1^2, \epsilon_1 D, \epsilon_2\} \|w\|_2^2$.

Proof. We have that

$$\text{Var}_{x \sim \mu}[w \cdot x] = \mathbf{E}_{x \sim \mu}[(w \cdot x)^2] - \left(\mathbf{E}_{x \sim \mu}[w \cdot x] \right)^2.$$

We first deal with upper-bounding the square of the first moment. Note that

$$w \cdot \left(\mathbf{E}_{x \sim \mu}[x] - \mathbf{E}_{x \sim U}[x] \right) \leq \|w\|_2 \left\| \mathbf{E}_{x \sim \mu}[x] - \mathbf{E}_{x \sim U}[x] \right\|_2 \leq \epsilon_1 \|w\|_2.$$

Let us take $\epsilon > 0$ (with $\epsilon < \epsilon_1$) for simplicity to be such that $(\mathbf{E}_{x \sim \mu}[w \cdot x])^2 = (\mathbf{E}_{x \sim U}[w \cdot x] + \epsilon \|w\|_2)^2$. This means that

$$\begin{aligned} \left(\mathbf{E}_{x \sim \mu}[w \cdot x] \right)^2 &\leq \left(\mathbf{E}_{x \sim U}[w \cdot x] \right)^2 + 2\epsilon \|w\|_2 \left| \mathbf{E}_{x \sim U}[w \cdot x] \right| + \epsilon^2 \|w\|_2^2 \\ &\leq \left(\mathbf{E}_{x \sim U}[w \cdot x] \right)^2 + 2\epsilon D \|w\|_2^2 + \epsilon^2 \|w\|_2^2. \end{aligned}$$

Next we lower-bound the second moment. It holds that

$$\mathbf{E}_{x \sim \mu}[(w \cdot x)^2] = w^\top \mathbf{E}_{x \sim \mu}[xx^\top]w \geq \mathbf{E}_{x \sim U}[(w \cdot x)^2] - \epsilon_2 \|w\|_2^2,$$

for some $\epsilon_2 > 0$. This means that

$$\text{Var}_{x \sim \mu}(w \cdot x) \geq \mathbf{E}_{x \sim U}[(w \cdot x)^2] - \epsilon_2 \|w\|_2^2 - \left(\mathbf{E}_{x \sim U}[w \cdot x] \right)^2 - 2\epsilon D \|w\|_2^2 - \epsilon^2 \|w\|_2^2.$$

Hence,

$$\text{Var}_{x \sim \mu}[w \cdot x] \geq \text{Var}_{x \sim U}[w \cdot x] - 3 \max\{\epsilon_2, \epsilon_1^2, \epsilon_1 D\} \|w\|_2^2.$$

□

Our next goal is to relate $\phi(\cdot; \rho^* w)$ with the uniform measure $\phi(\cdot; 0)$. According to the above general lemma, we have to relate the first and second moments of $\phi(\cdot; \rho^* w)$ with the ones of the uniform distribution $U = \phi(\cdot; 0)$.

The Proof of Proposition 8. Our goal is to apply Lemma 8. First, let us set

$$f_v(\rho) = \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[v \cdot x],$$

for any unit vector $v \in \Theta$. Then it holds that

$$\left\| \mathbf{E}_{x \sim \phi(\cdot; 0)}[x] - \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[x] \right\|_2 = \sup_{v: \|v\|_2=1} |f_v(0) - f_v(\rho^*)|.$$

Using the mean value theorem in $[0, \rho^*]$ for any unit vector v , we have that there exists a $\xi = \xi_v \in (0, \rho^*)$ such that

$$|f_v(0) - f_v(\rho^*)| = \rho^* |f'_v(\xi)|.$$

It suffices to upper bound $f'_v(\xi)$ for any unit vector v and $\xi \in (0, \rho^*)$. Let us compute f'_v . We have that

$$\frac{df_v}{d\rho} = \int_S (v \cdot x) \frac{d}{d\rho} \frac{e^{\rho(w \cdot x)}}{\int_S e^{\rho(w \cdot y)} dy} dx = \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(v \cdot x)(w \cdot x)] - \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[v \cdot x] \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[w \cdot x].$$

880 Since $x \in \mathcal{X}$, we have that

$$\sup_{v: \|v\|_2=1} \sup_{\xi \in (0, \rho^*)} |f'_v(\xi)| \leq 2\|w\|_2 D^2.$$

881 This gives that

$$\left\| \mathbf{E}_{x \sim \phi(\cdot; 0)}[x] - \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[x] \right\|_2 \leq 2\rho^* \|w\|_2 D^2.$$

882 We then continue with controlling the second moment: it suffices to find ϵ_2 such that for any $v \in \Theta$,
883 it holds

$$v^\top \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[xx^\top]v \geq v^\top \mathbf{E}_{x \sim \phi(\cdot; 0)}[xx^\top]v - \epsilon_2 \|v\|_2^2.$$

884 Let us set $g_v(\rho) = \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(v \cdot x)^2]$ for any vector $v \in \Theta$. We have that

$$|g_v(0) - g_v(\rho^*)| = \rho^* |g'_v(\xi)|,$$

885 where $\xi \in (0, \rho^*)$. It holds that

$$\left| \frac{dg_v}{d\rho} \right| = \left| \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(v \cdot x)^2(w \cdot x)] - \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[(v \cdot x)^2] \mathbf{E}_{x \sim \phi(\cdot; \rho w)}[w \cdot x] \right| \leq 2\|v\|^2 \|w\|_2 D^3.$$

886 This gives that for any $v \in \Theta$, it holds

$$v^\top \mathbf{E}_{x \sim \phi(\cdot; \rho^* w)}[xx^\top]v \geq v^\top \mathbf{E}_{x \sim \phi(\cdot; 0)}[xx^\top]v - 2\rho^* \|w\|_2 D^3 \|v\|_2^2.$$

887 Note that the above holds for $v = c$ too. [Lemma 8](#) gives us that

$$\mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[c \cdot x] \geq \mathbf{Var}_{x \sim \phi(\cdot; 0)}[c \cdot x] - 3 \max\{\epsilon_2, \epsilon_1^2, \epsilon_1 D\} \|c\|_2^2,$$

888 where $\epsilon_1 = 2\rho^* B D^2$ and $\epsilon_2 = 2\rho^* B D^3$. This implies that by picking

$$\rho^* = C_0 \frac{\alpha}{B D^3}$$

889 for some universal constant C_0 , we get that

$$\mathbf{Var}_{x \sim \phi(\cdot; \rho^* w)}[c \cdot x] = \Omega(\alpha \|c\|_2^2).$$

890

□

891 In this section, we considered ρ^* as indicated by the above [Proposition 8](#).

892 G Applications to Combinatorial Problems

893 In this section we provide a series of combinatorial applications of our theoretical framework
894 ([Theorem 1](#)). In particular, for each one of the following combinatorial problems (that provably
895 satisfy [Assumption 1](#)), it suffices to specify the feature mappings $\psi_S, \psi_{\mathcal{I}}$ and compute the parameters
896 $C, D_S, D_{\mathcal{I}}, \alpha$.

897 G.1 Maximum Cut, Maximum Flow and Max- k -CSPs

898 We first provide a general lemma for the variance of "linear tensors" under the uniform measure.

899 **Lemma 9** (Variance Lower Bound Under Uniform). *Let $n, k \in \mathbb{N}$. For any $w \in \mathbb{R}^{\binom{n}{k}}$, it holds that*

$$\mathbf{Var}_{x \sim U(\{-1, 1\}^n)}[w \cdot x^{\otimes k}] = \sum_{\emptyset \neq S \subseteq [n]: |S| \leq k} w_S^2.$$

900 *Proof.* For any $w \in \mathbb{R}^{\binom{n}{k}}$, it holds that

$$\mathbf{Var}_{x \sim U(\{-1, 1\}^n)}[w \cdot x^{\otimes k}] = \mathbf{E}_{x \sim U(\{-1, 1\}^n)}[(w \cdot x^{\otimes k})^2] - \mathbf{E}_{x \sim U(\{-1, 1\}^n)}[w \cdot x^{\otimes k}]^2.$$

901 Note that $w \in \mathbb{R}^{\binom{n}{k}}$ can be written as $w = (w_\emptyset, w_{-\emptyset})$ where w_\emptyset corresponds to the constant term of
902 the Fourier expansion and $w_{-\emptyset} = (w_S)_{\emptyset \neq S \subseteq [n]: |S| \leq k}$ is the vector of the remaining coordinates. The
903 Fourier expansion implies that

$$\mathbf{Var}_{x \sim U(\{-1, 1\}^n)}[w \cdot x^{\otimes k}] = \|w_{-\emptyset}\|_2^2,$$

904 which yields the desired equality for the variance. □

905 G.1.1 Maximum Cut

906 Let us consider a graph with n nodes and weighted adjacency matrix A with non-negative weights.
 907 Maximum cut is naturally associated with the Ising model and, intuitively, our approach does not
 908 yield an efficient algorithm for solving Max-Cut since we cannot efficiently sample from the Ising
 909 model in general. To provide some further intuition, consider a single-parameter Ising model for
 910 $G = (V, E)$ with Hamiltonian $H_G(x) = \sum_{(i,j) \in E} \frac{1+x_i x_j}{2}$. Then the partition function is equal
 911 to $Z_G(\beta) = \sum_{x \in \{-1,1\}^V} \exp(\beta H_G(x))$. Note that when $\beta > 0$, the Gibbs measure favours
 912 configurations with aligned spins (ferromagnetic case) and when $\beta < 0$, the measure favours
 913 configurations with opposite spins (anti-ferromagnetic case). The antiferromagnetic Ising model
 914 appears to be more challenging. According to physicists the main reason is that its Boltzmann
 915 distribution is prone to a complicated type of long-range correlation known as ‘replica symmetry
 916 breaking’ [COLMS22]. From the TCS viewpoint, observe that as β goes to $-\infty$, the mass of the
 917 Gibbs distribution shifts to spin configurations with more edges joining vertices with opposite spins
 918 and concentrates on the maximum cuts of the graph. Hence, being able to efficiently approximate the
 919 log-partition function for general Ising models, would lead to solving the Max-Cut problem.

920 **Theorem 2** (Max-Cut has a Compressed and Efficiently Optimizable Solution Generator). *Consider*
 921 *a prior over Max-Cut instances with n nodes. For any $\epsilon > 0$, there exists a solution generator*
 922 *$\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$ such that \mathcal{P} is complete, compressed with description $\text{poly}(n)\text{polylog}(1/\epsilon)$*
 923 *and $\mathcal{L} + \lambda R : \mathcal{W} \mapsto \mathbb{R}$ is efficiently optimizable via projected stochastic gradient descent in*
 924 *$\text{poly}(n, 1/\epsilon)$ steps for some $\lambda > 0$.*

925 *Proof of Theorem 2.* It suffices to show that Max-Cut satisfies Assumption 1. Consider an input
 926 graph G with n nodes and Laplacian matrix L_G . Then

$$\text{MAXCUT} = \frac{1}{4} \max_{s \in \{-1,1\}^n} s^\top L_G s = \frac{1}{4} \min_{s \in \{-1,1\}^n} -s^\top L_G s.$$

927 We show that there exist feature mappings so that the cost of every solution s under any instance/graph
 928 G is a bilinear function of the feature vectors (cf. Item 2 of Assumption 1). We consider the correlation-
 929 based feature mapping $\psi_S(s) = (ss^\top)^\flat \in \mathbb{R}^{n^2}$, where by $(\cdot)^\flat$ we denote the vectorization/flattening
 930 operation and the negative Laplacian for the instance (graph), $\psi_{\mathcal{I}}(G) = (-L_G)^\flat \in \mathbb{R}^{n^2}$. Then simply
 931 setting the matrix M to be the identity $I \in \mathbb{R}^{n^2 \times n^2}$ the cost of any solution s can be expressed as the
 932 bilinear function $\psi_{\mathcal{I}}(G)^\top M \psi_S(s) = (-L_G)^\top (ss^\top)^\flat = -s^\top L_G s$. We observe that (for unweighted
 933 graphs) with n nodes the bit-complexity of the family of all instances \mathcal{I} is roughly $O(n^2)$, and
 934 therefore the dimensions of the $\psi_S, \psi_{\mathcal{I}}$ feature mappings are clearly polynomial in the bit-complexity
 935 of \mathcal{I} . Moreover, considering unweighted graphs, it holds $\|\psi_{\mathcal{I}}(G)\|_2, \|\psi_S(s)\|_2, \|M\|_F \leq \text{poly}(n)$.
 936 Therefore, the constants $D_S, D_{\mathcal{I}}, C$ are polynomial in the bit-complexity of the instance family.

937 It remains to show that our solution feature mapping satisfy the variance preservation assumption.
 938 For any v , we have that $\text{Var}_{s \sim U(S)}[v \cdot \psi_S(s)] = \text{Var}_{s \sim U(\{-1,1\}^n)}[v \cdot (ss^\top)^\flat] = \Omega(\|v\|_2^2)$, using
 939 Lemma 9 with $k = 2$, since $c_\emptyset = 0$ with loss of generality. \square

940 G.1.2 Minimum Cut/Maximum Flow

941 Let us again consider a graph with n nodes and Laplacian matrix L_G . It is known that the minimum
 942 cut problem is solvable in polynomial time when all the weights are positive. From the discussion
 943 of the maximum cut case, we can intuitively relate minimum cut with positive weights to the
 944 ferromagnetic Ising setting [dPS97]. We remark that we can consider the ferromagnetic parameter
 945 space $\mathcal{W}_{\text{fer}} = \mathbb{R}_{\geq 0}^{\binom{n}{2}}$ and get the variance lower bound from Lemma 9. We constraint projected
 946 SGD in \mathcal{W}_{fer} . This means that during any step of SGD our algorithm has to sample from a mixture
 947 of ferromagnetic models with known mixture weights. The state of the art approximate sampling
 948 algorithm from ferromagnetic Ising models achieves the following performance, improving on prior
 949 work [JS93, LSS19, CLV22, CGG⁺19].

950 **Proposition 9** (Theorem 1.1 of [CZ22]). *Let $\delta_\beta, \delta_\lambda \in (0, 1)$ be constants and μ be the Gibbs*
 951 *distribution of the ferromagnetic Ising model specified by graph $G = (V, E), |V| = n, |E| = m$,*
 952 *parameters $\beta \in [1 + \delta_\beta, +\infty)^m$ and external field $\lambda \in [0, 1 - \delta_\lambda]^n$. There exists an algorithm that*

953 samples X satisfying $\text{TV}(X, \mu) \leq \epsilon$ for any given parameter $\epsilon \in (0, 1)$ within running time

$$m \left(\frac{\log n}{\epsilon} \right)^{O_{\delta_\beta, \delta_\lambda}(1)}.$$

954 This algorithm can handle general instances and it only takes a near-linear running time when
 955 parameters are bounded away from the all-ones vector. Our goal is to sample from a mixture of
 956 two such ferromagnetic Ising models which can be done efficiently. For simplicity, we next restrict
 957 ourselves to the unweighted case.

958 **Theorem 3** (Min-Cut has a Compressed, Efficiently Optimizable and Samplable Solution Generator).
 959 Consider a prior over Min-Cut instances with n nodes. For any $\epsilon > 0$, there exists a solution generator
 960 $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$ such that \mathcal{P} is complete, compressed with description $\text{poly}(n)\text{polylog}(1/\epsilon)$,
 961 $\mathcal{L} + \lambda R : \mathcal{W} \mapsto \mathbb{R}$ is efficiently optimizable via projected stochastic gradient descent in $\text{poly}(n, 1/\epsilon)$
 962 steps for some $\lambda > 0$ and efficiently samplable in $\text{poly}(n, 1/\epsilon)$ steps.

963 *Proof.* We have that

$$\text{MINCUT} = \frac{1}{4} \min_{x \in \{-1, 1\}^n} x^\top L_G x.$$

964 The analysis (i.e., the selection of the feature mappings) is similar to the one of [Theorem 2](#) with the
 965 sole difference that the parameter space is constrained to be \mathcal{W}_{fer} and $\psi_{\mathcal{I}}(G) = (L_G)^b$. We note that
 966 [Proposition 9](#) is applicable during the optimization steps. Having an efficient approximate sampler
 967 for solutions of Min-Cut, it holds that the runtime of the projected SGD algorithm is $\text{poly}(n, 1/\epsilon)$.
 968 We note that during the execution of the algorithm we do not have access to perfectly unbiased
 969 samples from mixture of ferromagnetic Ising models. However, we remark that SGD is robust to that
 970 inaccuracy in the stochastic oracle. For further details, we refer e.g., to [\[d'A08\]](#). \square

971 G.1.3 Max- k -CSPs

972 In this problem, we are given a set of variables $\{x_u\}_{u \in \mathcal{U}}$ where $|\mathcal{U}| = n$ and a set of Boolean
 973 predicates P . Each variable x_u takes values in $\{-1, 1\}$. Each predicate depends on at most k
 974 variables. For instance, Max-Cut is a Max-2-CSP. Our goal is to assign values to variables so as to
 975 maximize the number of satisfied constraints (i.e., predicates equal to 1). Let us fix a predicate $h \in P$,
 976 i.e., a Boolean function $h : \{-1, 1\}^n \rightarrow \{0, 1\}$ which is a k -junta. Using standard Fourier analysis,
 977 the number of satisfied predicates for the assignment $x \in \{-1, 1\}^n$ is

$$F(x) = \sum_{j=1}^{|P|} \sum_{S \subseteq [n], |S| \leq k} \hat{h}_j(S) \prod_{u \in S} x_u,$$

978 where $\hat{h}_j(S)$ is the Fourier coefficient of the predicate h_j at S .

979 **Theorem 4** (Max- k -CSPs have a Compressed and Efficiently Optimizable Solution Generator).
 980 Consider a prior over Max- k -CSP instances with n variables, where $k \in \mathbb{N}$ can be considered
 981 constant compared to n . For any $\epsilon > 0$, there exists a solution generator $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$
 982 such that \mathcal{P} is complete, compressed with description $O(n^k)\text{polylog}(1/\epsilon)$ and $\mathcal{L} + \lambda R : \mathcal{W} \mapsto \mathbb{R}$
 983 is efficiently optimizable via projected stochastic gradient descent in $\text{poly}(n^k, 1/\epsilon)$ steps for some
 984 $\lambda > 0$.

985 *Proof.* Any instance of Max- k -CSP is a list of predicates (i.e., Boolean functions) and our goal
 986 is to maximize the number of satisfied predicated with a single assignment $s \in \{-1, 1\}^n$. We
 987 show that there exist feature mappings so that the cost of every solution s under any instance/predi-
 988 cates list P is a bilinear function of the feature vectors (cf. Item 2 of [Assumption 1](#)). We con-
 989 sider the order k correlation-based feature mappings $\psi_S(s) = (s^{\otimes k})^b \in \mathbb{R}^{n^k}$, where by $(\cdot)^b$
 990 we denote the flattening operation of the order k tensor, and, $\psi_{\mathcal{I}}(P) = \psi_{\mathcal{I}}(h_1, \dots, h_{|P|}) =$
 991 $-\sum_{j=1}^{|P|} ((\hat{h}_j(S))_{S \subseteq [n], |S| \leq k})^\top \in \mathbb{R}^{n^k}$, where $(\hat{h}_j(S))_{S \subseteq [n], |S| \leq k}$ is a vector of size n^k with
 992 the Fourier coefficients of the j -th predicate. We take $\psi_{\mathcal{I}}$ being the coordinate-wise sum of
 993 these coefficients. The setting the matrix M to be the identity matrix $I \in \mathbb{R}^{n^k \times n^k}$, we get
 994 that the cost of any solution s can be expressed as the bilinear function $\psi_{\mathcal{I}}(P)^\top M \psi_S(s) =$

995 $-\sum_{j=1}^{|P|} \sum_{S \subseteq [n], |S| \leq k} \hat{h}_j(S) \prod_{u \in S} x_u$. For any $h : \{-1, 1\}^n \rightarrow \{0, 1\}$, we get that the description
 996 size of any $\hat{h}(S)$ is $\text{poly}(n, k)$ and so the dimensions of the $\psi_S, \psi_{\mathcal{I}}$ feature mappings are polynomial
 997 in the description size of \mathcal{I} . Moreover, we get that $\|\psi_{\mathcal{I}}(P)\|, \|\psi_S(s)\|, \|M\| \leq \text{poly}(n^k)$. Hence,
 998 the constants $D_S, D_{\mathcal{I}}, C$ are polynomial in the description size of the instance family. Finally, we
 999 have that for any v , $\text{Var}_{s \sim U(S)}[v \cdot \psi_S(s)] = \text{Var}_{s \sim U(\{-1, 1\}^n)}[v \cdot s^{\otimes k}] = \Omega(\|v\|_2^2)$, using [Lemma 9](#),
 1000 assuming that v_{\emptyset} is 0 without loss of generality. This implies the result. \square

1001 G.2 Bipartite Matching and TSP

1002 G.2.1 Maximum Weight Bipartite Matching

1003 In Maximum Weight Bipartite Matching (MWBM) there exists a complete bipartite graph (A, B)
 1004 with $|A| = |B| = n$ (the assumptions that the graph is complete and balanced is without loss of
 1005 generality) with weight matrix W where $W(i, j)$ indicates the value of the edge (i, j) , $i \in A, j \in B$
 1006 and the goal is to match the vertices in order to maximize the value. Hence the goal is to maximize
 1007 $L(\Pi) = W \cdot \Pi$ over all permutation matrices. By the structure of the problem some maximum
 1008 weight matching is a perfect matching. Furthermore, by negating the weights of the edges we
 1009 can state the problem as the following minimization problem: given a bipartite graph (A, B) and
 1010 weight matrix $W \in (\mathbb{R} \cup \{\infty\})^{n \times n}$, find a perfect matching M with minimum weight. One of the
 1011 fundamental results in combinatorial optimization is the polynomial-time blossom algorithm for
 1012 computing minimum-weight perfect matchings by [[Edm65](#)].

1013 We begin this section by showing a variance lower bound under the uniform distribution over the
 1014 permutation group.

1015 **Lemma 10** (Variance Lower Bound). *Let $U(\mathbb{S}_n)$ be the uniform distribution over $n \times n$ permutation*
 1016 *matrices. For any matrix $W \in \mathbb{R}^{n \times n}$, with $\sum_i W_{ij} = 0$ and $\sum_j W_{ij} = 0$ we have*

$$\text{Var}_{\Pi \sim U(\mathbb{S}_n)}[W \cdot \Pi] = \frac{\|W\|_F^2}{n-1}.$$

1017 *Proof.* We have that $\mathbf{E}_{\Pi \sim U(\mathbb{S}_n)}[\Pi_{ij}] = 1/n$ and $\mathbf{E}_{\Pi \sim U(\mathbb{S}_n)}[\Pi_{ij}\Pi_{ab}] = \frac{1\{i \neq a, j \neq b\}}{n(n-1)} + \frac{1\{i=a, j=b\}}{n}$.
 1018 We have

$$\begin{aligned} \text{Var}_{\Pi \sim U(\mathbb{S}_n)}[W \cdot \Pi] &= \mathbf{E}_{\Pi \sim U(\mathbb{S}_n)}[(W \cdot \Pi)^2] - \left(\mathbf{E}_{\Pi \sim U(\mathbb{S}_n)}[W \cdot \Pi] \right)^2 \\ &= \sum_{i,j,a,b} W_{ab} W_{ij} \left(\frac{1\{i \neq a, j \neq b\}}{n(n-1)} + \frac{1\{i=a, j=b\}}{n} \right) - \left(\sum_{i,j} \frac{W_{ij}}{n} \right)^2 \\ &= \frac{1}{n} \sum_{i,j} W_{ij}^2 + \sum_{i,j,a,b} W_{ij} W_{ab} \frac{1\{i \neq a, j \neq b\}}{n(n-1)} \\ &= \frac{\|W\|_F^2}{n} + \sum_{i,j,a,b} W_{ij} W_{ab} \frac{1\{i \neq a, j \neq b\}}{n(n-1)}, \end{aligned}$$

1019 where to obtain the third equality we used our assumption that $\sum_{ij} W_{ij} = 0$. We observe that, by
 1020 our assumption that $\sum_b W_{ab} = 0$ for all a it holds $\sum_{b \neq j} W_{ab} = -W_{aj}$ and therefore, we have

$$\sum_b W_{ab} 1\{i \neq a, j \neq b\} = 1\{i \neq a\} \sum_b W_{ab} 1\{j \neq b\} = -1\{i \neq a\} W_{aj}.$$

1021 Similarly, using the fact that $\sum_{a \neq i} W_{aj} = -W_{ij}$ we obtain that

$$\sum_{ab} W_{ab} 1\{i \neq a, j \neq b\} = \sum_a -W_{aj} 1\{i \neq a\} = W_{ij}.$$

1022 Therefore, using the above identity, we have that

$$\sum_{i,j,a,b} W_{ij} W_{ab} \frac{1\{i \neq a, j \neq b\}}{n(n-1)} = \sum_{i,j} \frac{W_{ij}^2}{n(n-1)} = \frac{\|W\|_F^2}{n(n-1)}.$$

1023 Combining the above we obtain the claimed identity. \square

1024 **Remark 8.** We note that in MWBM the conditions $\sum_i W_{ij} = 0$ and $\sum_j W_{ij} = 0$ are without loss of
 1025 generality.

1026 We next claim that there exists an efficient algorithm for (approximately) sampling such permutation
 1027 matrices.

1028 **Lemma 11** (Efficient Sampling). *There exists an algorithm that generates approximate samples
 1029 from the Gibbs distribution $p(\cdot; W)$ with parameter W over the symmetric group, i.e., $p(\Pi; W) \propto$
 1030 $\exp(W \cdot \Pi) 1\{\Pi \in \mathbb{S}_n\}$, in $\text{poly}(n)$ time.*

1031 *Proof.* This lemma essentially requires approximating the permanent of a weighted matrix, since
 1032 this would imply that one has an approximation of the partition function. Essentially, our goal is to
 1033 generate a random variable X that is ϵ -close in statistical distance to the probability measure

$$p(\Pi; W) \propto \exp(W \cdot \Pi) 1\{\Pi \in \mathbb{S}_n\}.$$

1034 Note that the partition function is

$$Z(W) = \sum_{\Pi \in \mathbb{S}_n} e^{W \cdot \Pi} = \sum_{\Pi \in \mathbb{S}_n} \prod_{(i,j)} e^{W_{ij} \Pi_{ij}} = \sum_{\sigma \in \mathbb{S}_n} \prod_{i \in [n]} A_{i, \sigma(i)},$$

1035 where A is a non-negative real matrix with entries $A_{ij} = \exp(W_{ij})$. Note that in the third equality,
 1036 we used the isomorphism between permutations and permutation matrices. Hence, $Z(W)$ is exactly
 1037 the permanent of the matrix A .

1038 **Proposition 10** ([JSV04]). *There exists a fully polynomial randomized approximation scheme for the
 1039 permanent of an arbitrary $n \times n$ matrix A with non-negative entries.*

1040 To conclude the proof of the lemma, we need the following standard result.

1041 **Proposition 11** (See [Appendix H](#) and [Sin12, Jer03]). *For self-reducible problems, fully polynomial
 1042 approximate integration and fully polynomial approximate sampling are equivalent.*

1043 This concludes the proof since weighted matchings are self-reducible (see [Appendix H](#)). \square

1044 The above lemma establishes our goal:

1045 **Theorem 5** (MWBM has a Compressed, Efficiently Optimizable and Samplable Solution Gen-
 1046 erator). *Consider a prior over MWBM instances with n nodes. For any $\epsilon > 0$, there exists a
 1047 solution generator $\mathcal{P} = \{p(w) : w \in \mathcal{W}\}$ such that \mathcal{P} is complete, compressed with description
 1048 $\text{poly}(n)\text{polylog}(1/\epsilon)$, $\mathcal{L} + \lambda R : \mathcal{W} \mapsto \mathbb{R}$ is efficiently optimizable via projected stochastic gradient
 1049 descent in $\text{poly}(n, 1/\epsilon)$ steps for some $\lambda > 0$ and efficiently samplable in $\text{poly}(n, 1/\epsilon)$ steps.*

1050 *Proof.* Consider an input graph G with n nodes and adjacency matrix E . The feature vector
 1051 corresponding to a matching can be represented as a binary matrix $\Pi \in \{0, 1\}^{n \times n}$ with $\sum_j \Pi_{ij} = 1$
 1052 for all i and $\sum_i \Pi_{ij} = 1$ for all j , i.e., Π is a permutation matrix. Then

$$\text{MWBM} = \max_{\Pi \in \mathbb{S}_n} E \cdot \Pi = \min_{\Pi \in \mathbb{S}_n} -E \cdot \Pi.$$

1053 Therefore, for a candidate matching s , we set $\psi_S(s)$ to be the matrix Π defined above. Moreover,
 1054 the feature vector of the graph is the negative (flattened) adjacency matrix $-E^b$. The cost oracle
 1055 is then $L(R; E) = -\sum_{ij} E_{ij} M_{ij} R_{ij}$ perhaps for an unknown weight matrix M_{ij} (see [Remark 6](#)).
 1056 This means that the dimensions of the feature mappings $\psi_S, \psi_{\mathcal{I}}$ are polynomial in the bit complexity
 1057 of \mathcal{I} . Moreover, we get that $\|\psi_{\mathcal{I}}(I)\|_F, \|\psi_S(s)\|_F, \|M\|_F \leq \text{poly}(n)$. We can employ [Lemma 10](#)
 1058 to get the variance lower bound under the uniform probability distribution in the subspace induced
 1059 by the matrices satisfying [Lemma 10](#), i.e., the matrices of the parameter space (see [Remark 9](#)).
 1060 Finally, (approximate) sampling from our solution generators can be done efficiently using [Lemma 11](#)
 1061 and hence (noisy) projected SGD will have a runtime of order $\text{poly}(n, 1/\epsilon)$ (as in the case of Min-
 1062 Cut). \square

1063 We close this section with a remark about Item 3 of [Assumption 1](#).

Remark 9. We note that Item 3 of [Assumption 1](#) can be weakened. We use our variance lower bound in order to handle inner products of the form $w \cdot x$ where w will lie in the parameter space and x is the featurization of a solution that lies in some space X . Hence it is possible that w lies in a low-dimensional subspace of X . For our optimization purposes, it suffices to provide variance lower bounds only in the subspace where w lies into.

G.2.2 Travelling Salesman Problem

Let us consider a weighted clique K_n with n vertices and weight matrix $W \in \mathbb{R}^{n \times n}$. A solution to the TSP instance W is a sequence $\pi : [n] \rightarrow [n]$ of the n elements indicating the TSP tour $(\pi(1), \pi(2), \dots, \pi(n), \pi(1))$ and suffers a cost

$$L(\pi) = \sum_{i=1}^{n-1} W_{\pi(i), \pi(i+1)} + W_{\pi(n), \pi(1)}.$$

Crucially, the allowed sequences are a proper subset of all possible permutations. For instance, the permutations with fixed points or small cycles are not allowed. In particular, the solution space of TSP corresponds to the set of cyclic permutations with no trivial cycles, i.e., containing an n -cycle. Clearly, the number of n -cycles is $(n-1)!$. The goal is to find a tour of minimum cost. Our first goal is to write the cost objective as a linear function of the weight matrix W and the feasible solutions, which correspond to cyclic permutations. To this end, we can think of each cyclic permutation π as a cyclic permutation matrix $\Pi \in \{0, 1\}^{n \times n}$. Then, the desired linearization is given by $L(\Pi) = W \cdot \Pi$ (for a fixed graph instance).

Our next task is to provide a Gibbs measure that generates random cyclic permutations. Let \mathbb{C}_n be the space of $n \times n$ cyclic permutation matrices. Then we have that the tour Π is drawn from

$$p_W(\Pi) = \frac{\exp(W \cdot \Pi) 1\{\Pi \in \mathbb{C}_n\}}{\sum_{\Pi' \in \mathbb{C}_n} \exp(W \cdot \Pi')} ,$$

where W is the weight matrix. The following key lemma provides guarantees for the performance of our approach to TSP. This lemma allows us to show that the number of optimization steps is $\text{poly}(n, 1/\epsilon)$.

Lemma 12 (Variance Lower Bound). *Let $U(\mathbb{C}_n)$ be the uniform distribution over $n \times n$ cyclic permutation matrices. For any matrix $W \in \mathbb{R}^{n \times n}$, with $\sum_i W_{ij} = 0$ and $\sum_j W_{ij} = 0$ we have*

$$\text{Var}_{\Pi \sim U(\mathbb{C}_n)}[W \cdot \Pi] \geq \frac{\|W\|_F^2}{(n-1)(n-2)}.$$

Proof. The first step is to compute some standard statistics about cyclic permutations (see [Lemma 13](#)). [Lemma 13](#) and the analysis of [Lemma 10](#) gives us that $\text{Var}_{\Pi \sim U(\mathbb{C}_n)}[W \cdot \Pi]$ is equal to

$$\frac{\|W\|_F^2}{n-1} + \sum_{i,j,a,b} W_{ij} W_{ab} \left(\frac{1\{i \neq a = j \neq b\}}{(n-1)(n-2)} + \frac{1\{j \neq b = i \neq a\}}{(n-1)(n-2)} + \frac{1\{i \neq a \neq j \neq b \neq i\}}{(n-1)(n-3)} \right).$$

Let us set

$$A_1 = \sum_{i,j,a,b} W_{ij} W_{ab} 1\{i \neq a = j \neq b\}.$$

We have that

$$\sum_b W_{ab} 1\{i \neq a\} 1\{a = j\} 1\{j \neq b\} = 1\{i \neq a\} 1\{a = j\} \sum_b W_{ab} 1\{j \neq b\} = -1\{i \neq a\} 1\{a = j\} W_{aj}.$$

Hence

$$A_1 = - \sum_{i,j,a} W_{ij} W_{aj} 1\{i \neq a\} 1\{a = j\} = - \sum_{i,j} W_{ij} W_{jj} 1\{i \neq j\} = \sum_j W_{jj}^2.$$

Due to symmetry, $A_2 = \sum_{i,j,a,b} W_{ij} W_{ab} 1\{j \neq b = i \neq a\} = \sum_j W_{jj}^2$. It remains to argue about

$$A_3 = \sum_{i,j,a,b} W_{ij} W_{ab} 1\{i \neq a \neq j \neq b \neq i\}.$$

1094 We have that

$$\sum_b W_{ab} 1\{i \neq a\} 1\{a \neq j\} 1\{j \neq b\} 1\{b \neq i\} = -1\{i \neq a\} 1\{a \neq j\} (W_{aj} + W_{ai}).$$

1095 This gives that

$$A_3 = - \sum_{i,j,a} W_{ij} (W_{aj} + W_{ai}) 1\{i \neq a\} 1\{a \neq j\}.$$

1096 Note that

$$\sum_{i,j,a} W_{ij} W_{aj} 1\{i \neq a\} 1\{a \neq j\} = \sum_{j,a} W_{aj} 1\{a \neq j\} \sum_i W_{ij} 1\{i \neq a\} = - \sum_{j,a} W_{aj}^2 1\{a \neq j\}.$$

1097 This implies that

$$A_3 = \sum_{j \neq a} W_{aj}^2 + \sum_{i \neq a} W_{ai}^2 = 2 \sum_{j \neq a} W_{aj}^2.$$

1098 In total, this gives that

$$\text{Var}_{\Pi \sim U(\mathbb{C}_n)}[W \cdot \Pi] = \frac{\|W\|_F^2}{n-1} + \frac{2 \sum_j W_{jj}^2}{(n-1)(n-2)} + \frac{2 \sum_{i \neq j} W_{ij}^2}{(n-1)(n-3)} \geq \frac{\|W\|_F^2}{n-1} + \frac{\|W\|_F^2}{(n-1)(n-2)}.$$

1099

□

1100 **Remark 10.** We note that in TSP the conditions $\sum_i W_{ij} = 0$ and $\sum_j W_{ij} = 0$ are without loss of
1101 generality.

1102 The next lemma is a generic lemma that states some properties of random cyclic permutation matrices.

1103 **Lemma 13.** Consider a uniformly random cyclic permutation matrix Π . Let us fix $i \neq j$ and $a \neq b$.
1104 Then

$$\begin{aligned} 1105 & \bullet \mathbf{E}[\Pi_{ij}] = \frac{1}{n-1}. \\ 1106 & \bullet \mathbf{E}[\Pi_{ij} \Pi_{ab}] = \frac{1\{i \neq a=j \neq b\}}{(n-1)(n-2)} + \frac{1\{j \neq b=i \neq a\}}{(n-1)(n-2)} + \frac{1\{i \neq a \neq j \neq b \neq i\}}{(n-1)(n-3)} + \frac{1\{i=a, j=b\}}{n-1}. \end{aligned}$$

1107 *Proof.* First, note that any matrix that corresponds to a cyclic permutation does not contain fixed
1108 points and so the diagonal elements are 0 deterministically. For the first item, the number of cyclic
1109 permutations such that $i \rightarrow j$ (i.e., $\Pi_{ij} = 1$) is $(n-2)!$. This implies that the desired expectation is
1110 $(n-2)!/|\mathbb{C}_n| = 1/(n-1)$. For the second item, if $i = a, j = b$, we recover the first item. Otherwise
1111 if $i = a$ or $j = b$, then the expectation vanishes since we deal with permutation matrices. Finally, let
1112 us consider the case where $i \neq a$ and $j \neq b$. Our goal is to count the number of cyclic permutations
1113 with $i \rightarrow j$ and $a \rightarrow b$.

- 1114 • If $i \neq a \neq j \neq b \neq i$, then there are n choices to place i and $n-2$ choices to place a . Then
1115 there are $(n-4)!$ possible orderings for the remaining elements. This gives an expectation
1116 equal to $1/((n-1)(n-3))$.
- 1117 • If $i \neq a = j \neq b$ or $j \neq b = i \neq a$, then there are n choices for i and $(n-3)!$ orderings for
1118 the remaining elements. Hence, the expectation is $1/((n-1)(n-2))$.

1119

□

1120 We note that sampling from our solution generators is the reason that we cannot find an optimal TSP
1121 solution efficiently. In general, an algorithm that has converged to an almost optimal parameter W^*
1122 has to generate samples from the Gibbs measure that is concentrated on cycles with minimum weight.
1123 In this low-temperature regime, sampling is NP-hard. We are now ready to state our result.

1124 **Theorem 6** (TSP has a Compressed, Efficiently Optimizable Solution Generator). Consider a
1125 prior over TSP instances with n nodes. For any $\epsilon > 0$, there exists a solution generator $\mathcal{P} =$
1126 $\{p(w) : w \in \mathcal{W}\}$ such that \mathcal{P} is complete, compressed with description $\text{poly}(n) \text{polylog}(1/\epsilon)$ and
1127 $\mathcal{L} + \lambda R : \mathcal{W} \mapsto \mathbb{R}$ is efficiently optimizable via projected stochastic gradient descent in $\text{poly}(n, 1/\epsilon)$
1128 steps for some $\lambda > 0$.

1129 *Proof.* Consider an input graph G with n nodes and weighted adjacency matrix E . The feature vector
 1130 is again a permutation matrix Π with the additional constraint that Π has to represent a single cycle
 1131 (a tour over all cities). Then

$$\text{TSP} = \min_{\Pi \in \mathbb{C}_n} E \cdot \Pi.$$

1132 The cost function for TSP is $L(\Pi; E) = \sum_{ij} E_{ij} M_{ij} \Pi_{ij}$. We refer to [Theorem 5](#) for the details
 1133 about the feature mappings. We can finally use [Lemma 12](#) to obtain a variance lower bound (in the
 1134 subspace induced by the parameters satisfying this lemma, see [Remark 9](#)) under the uniform measure
 1135 over the space of cyclic permutation matrices \mathbb{C}_n . \square

1136 H Sampling and Counting

1137 In this section, we give a quick overview of the connections between approximate sampling and
 1138 counting. For a formal treatment, we refer to [\[Sin12\]](#).

1139 In what follows, σ may be thought of as an encoding of an instance of some combinatorial problem,
 1140 and the ω of interest are encodings of the structures we wish to generate. Consider a weight function
 1141 W and assume that $W(\sigma, \omega)$ is computable in time polynomial in $|\sigma|$.

1142 **Definition 4** (Approximate Sampling). *A fully polynomial approximate sampler for $(\Omega_\sigma, \pi_\sigma)$ is*
 1143 *a Probabilistic Turing Machine which, on inputs σ and $\epsilon \in \mathbb{Q}_+$ ($0 < \epsilon \leq 1$), outputs $\omega \in \Sigma^*$,*
 1144 *according to a measure μ_σ satisfying $\text{TV}(\pi_\sigma, \mu_\sigma) \leq \epsilon$, in time bounded by a bivariate polynomial in*
 1145 *$|\sigma|$ and $\log(1/\epsilon)$.*

1146 One of the main applications of sampling is to approximate integration. In our setting this means
 1147 estimating $Z(\sigma)$ to some specified relative error.

1148 **Definition 5** (Approximate Integration). *A fully polynomial randomized approximation scheme for*
 1149 *$Z(\sigma)$ is a Probabilistic Turing Machine which on input σ, ϵ , outputs an estimate \hat{Z} so that*

$$\Pr[Z/(1 + \epsilon) \leq \hat{Z} \leq (1 + \epsilon)Z] \geq 3/4,$$

1150 *and which runs in time polynomial in $|\sigma|$ and $1/\epsilon$.*

1151 **Definition 6** (Self-Reducible Problems). *An NP search problem is self-reducible if the set of solutions*
 1152 *can be partitioned into polynomially many sets each of which is in a one-to-one correspondence*
 1153 *with the set of solutions of a smaller instance of the problem, and the polynomial size set of smaller*
 1154 *instances are efficiently computable.*

1155 For instance, consider the relation MATCH which associates with an undirected graph G all matchings
 1156 (independent sets of edges) of G . Then MATCH is self-reducible since, for any edge $e = (u, v) \in E(G)$,
 1157 we have that

$$\text{MATCH}(G) = \text{MATCH}(G_1) \cup \{M \cup \{e\} : M \in \text{MATCH}(G_2)\},$$

1158 where G_1 is the graph obtained by deleting e and G_2 is the graph obtained by deleting both u and v
 1159 together with all their incident edges.

1160 **Theorem 7** (See Corollary 3.16 in [\[Sin12\]](#)). *For self-reducible problems, approximate integration*
 1161 *and good sampling are equivalent.*

1162 We remark that the above result holds for the more general class of self-partitionable problems.

1163 I Details of the Experimental Evaluation

1164 We investigate the effect of the entropy regularizer (see [Equation \(2\)](#)) in a very simple setting: we try
 1165 to find the Max-Cut of a fixed graph G , i.e., the support of the prior \mathcal{R} is a single graph. We show that
 1166 while the unregularized objective is often “stuck” at sub-optimal solutions – and this happens even for
 1167 very small instances (15 nodes) – of the Max-Cut problem, the regularized version (with the fast/slow
 1168 mixture scheme) is able to find the optimal solutions. We consider an instance randomly generated
 1169 by the Erdős–Rényi model $G(n, p)$ and then optimize the “vanilla” loss \mathcal{L} and the regularized loss
 1170 \mathcal{L}_λ defined in [Equation \(3\)](#). The solutions are vectors $s \in \{\pm 1\}^n$. We first use the feature mapping
 1171 $\psi_S(s) = (ss^\top)^b$ described in [Section 1.1](#) and an exponential family solution generator that samples a

```

class FastSlowMixture(torch.nn.Module):
def __init__(self, dimension, rho):
    """
    The Model parameters.
    """
    super().__init__()

    self.l1 = torch.nn.Parameter(torch.empty(30, dimension))
    torch.nn.init.kaiming_uniform_(self.l1, a=5**0.5)

    self.l2 = torch.nn.Parameter(torch.empty(10, 30))
    torch.nn.init.kaiming_uniform_(self.l2, a=5**0.5)

    self.l3 = torch.nn.Parameter(torch.empty(1, 10))
    torch.nn.init.kaiming_uniform_(self.l3, a=5**0.5)

    self.a2 = torch.nn.ReLU()
    self.a1 = torch.nn.ReLU()

    self.rho = rho

def forward(self, x, is_cold=True):

    temp = self.rho * (1. - is_cold) + is_cold

    out = x
    out = torch.nn.functional.linear(out, temp * self.l1)
    out = self.a1(out)
    out = torch.nn.functional.linear(out, temp * self.l2)
    out = self.a2(out)
    out = torch.nn.functional.linear(out, temp * self.l3)

    return out

```

Figure 3: Our implementation of the fast/slow network. The output of the network is the log-density (score) of a solution $s \in \{\pm 1\}^{\text{dimension}}$. If evaluated with the is-cold set to False, the parameters of every linear layer are re-scaled by the inverse temperature rho.

1172 solution s with probability $\propto \exp(w \cdot \psi_S(s))$ for some weight vector $w \in \mathbb{R}^{n^2}$. We also consider
1173 optimizing a simple 3-layer ReLU network as solution generator with input $s \in \{\pm 1\}^n$ on the same
1174 random graphs. We generate 100 random $G(n, p)$ graphs with $n = 15$ nodes and $p = 0.5$ and train
1175 solution generators using both the "vanilla" and the entropy-regularized loss functions. We perform
1176 600 iterations and, for the entropy regularization, we progressively decrease the regularization weight,
1177 starting from 10, and dividing it by 2 every 60 iterations. We used a fast/slow mixing with mixture
1178 probability 0.2 and inverse temperature rho=0.03 (see Figure 3).

1179 For convenience, we present a pytorch implementation of our simple 3-layer ReLU network here.
1180 For more details we refer to our full code submitted in the supplementary material.

1181 Out of the 100 trials we found that our proposed objective was always able to find the optimal cut
1182 while the model trained with the vanilla loss was able to find it for approximately 65% of the graphs
1183 (for 65 out of 100 using the linear network and for 66 using the ReLU network). In Figure 2 we show
1184 two instances where the model trained with the "vanilla" loss gets stuck on a sub-optimal solution
1185 while the entropy-regularized one succeeds in finding the optimal solution. Our experiments show
1186 that the regularization term and the fast/slow mixture scheme that we introduced to achieve our main
1187 theoretical convergence result, see Section 3 and Proposition 4, are potentially useful for training
1188 more realistic models for bigger instances and we leave more extensive experimental evaluation as an
1189 interesting direction for future work.

1190 We note that, similarly to our theoretical results, our sampler in this experimental section is of the
 1191 form $e^{\text{score}(s;w)}$, where $s \in \{-1, 1\}^n$ (here n is the number of nodes in the graph) is a candidate
 1192 solution of the Max-Cut problem. The function used is a 3-layer MLP (see Figure [Figure 3](#)). Since
 1193 the instances that we consider here are small ($n = 15$) we can explicitly compute the density (score)
 1194 of every solution and use that to compute the expected gradient. For larger instances, one could use
 1195 some approximate sampler (e.g., via Langevin dynamics) to generate samples. The main message
 1196 of the current experimental section is that even for very small instances of Max-Cut (i.e., with 15
 1197 nodes), optimizing the vanilla objective is not sufficient and the iteration gets trapped in local optima.
 1198 In contrast, our entropy regularized always manages to find the optimal cut.