

Gaussian-Smoothed Sliced Probability Divergences: Supplementary Materials

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A Proofs

In the following sections, we give the proofs of the theoretical guarantees given in the main of the paper.

A.1 Proof of Theorem 3.1: $G_\sigma SD_p$ is a proper metric on $\mathcal{P}_p(\mathbb{R}^d) \times \mathcal{P}_p(\mathbb{R}^d)$

Before starting the proof, we add this notation: the characteristic function of a probability distribution $\mu \in \mathcal{P}(\mathbb{R}^d)$ is $\varphi_\mu(t) = \mathbf{E}_\mu[e^{iX^\top t}]$. Given this definition, similarly to the Fourier transform, the characteristic function of the convolution of two probability distributions reads as $\varphi_{\nu * \mu}(t) = \varphi_\nu(t) \cdot \varphi_\mu(t)$.

- *Non-negativity (or symmetry).* The non-negativity (or symmetry) follows directly from the non-negativity (or symmetry) of D^p , see Definition 2.3.

- *Identity property.* If the base divergence D^p satisfies the identity property in one dimensional measures, then for any $\mu \in \mathcal{P}_p(\mathbb{R}^d)$ and $\mathbf{u} \in \mathbb{S}^{d-1}$, one has that $D_p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) = 0$, hence, by Definition 2.3, $G_\sigma SD_p(\mu, \mu) = 0$. Let us now prove the fact that for any $\mu, \nu \in \mathcal{P}_p(\mathbb{R}^d)$, $G_\sigma SD_p(\mu, \nu) = 0$ entails $\mu = \nu$ a.s. On one hand, $G_\sigma SD_p(\mu, \nu) = 0$ gives the fact that $D_p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) = 0$ for u_d -almost every $\mathbf{u} \in \mathbb{S}^{d-1}$, hence $\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma = \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma$ for u_d -almost every $\mathbf{u} \in \mathbb{S}^{d-1}$. Following the techniques in proof of Proposition 5.1.2 in Bonnotte (2013), for any measure $\eta \in \mathcal{P}(\mathbb{R}^m)$ (with $m \geq 1$), $\mathcal{F}[\eta](\cdot)$ stands for the Fourier transform of η and is given as $\mathcal{F}[\eta](\mathbf{v}) = \int_{\mathbb{R}^m} e^{-i\mathbf{s}^\top \mathbf{v}} d\eta(\mathbf{s})$ for any $\mathbf{v} \in \mathbb{R}^m$. Then

$$\begin{aligned}
 \mathcal{F}[\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma](v) &= \int_{\mathbb{R}} e^{-ivt} d(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma)(t) \\
 &= \int_{\mathbb{R}} \int_{\mathbb{R}} e^{-i(r+t)v} d\mathcal{R}_{\mathbf{u}}\mu(r) d\mathcal{N}_\sigma(t) \quad (\text{by the definition of the convolution operator}) \\
 &= \int_{\mathbb{R}^d} \int_{\mathbb{R}} e^{-i(\langle \mathbf{u}, \mathbf{s} \rangle + t)v} d\mu(\mathbf{s}) d\mathcal{N}_\sigma(t) \quad (\text{by the definition of Radon Transform}) \\
 &= \int_{\mathbb{R}} e^{-itv} d\mathcal{N}_\sigma(t) \int_{\mathbb{R}^d} e^{-i(\langle \mathbf{u}, \mathbf{s} \rangle)v} d\mu(\mathbf{s}) \\
 &= \mathcal{F}[\mathcal{N}_\sigma](v) \mathcal{F}[\mu](v\mathbf{u}).
 \end{aligned}$$

Since for u_d -almost every $\mathbf{u} \in \mathbb{S}^{d-1}$, $\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma = \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma$, and hence $\mathcal{F}[\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma] = \mathcal{F}[\mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma] \Leftrightarrow \mathcal{F}[\mathcal{N}_\sigma]\mathcal{F}[\mu] = \mathcal{F}[\mathcal{N}_\sigma]\mathcal{F}[\nu]$ (by the Fourier transform of the convolution) $\Leftrightarrow \mathcal{F}[\mu] = \mathcal{F}[\nu]$. Since the Fourier

transform is injective, we conclude that $\mu = \nu$.

• *Triangle inequality.* Assume that D is a metric and let $\mu, \nu, \eta \in \mathcal{P}_p(\mathbb{R}^d)$. We then have

$$\begin{aligned} G_\sigma \text{SD}_p(\mu, \nu) &= \left(\int_{\mathbb{S}^{d-1}} D^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} \\ &\leq \left(\int_{\mathbb{S}^{d-1}} (D(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\eta * \mathcal{N}_\sigma) + D(\mathcal{R}_{\mathbf{u}}\eta * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma))^p u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} \\ &\stackrel{(\star)}{\leq} \underbrace{\left(\int_{\mathbb{S}^{d-1}} (D^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\eta * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u}) \right)^{1/p}}_{(\star)} + \left(\int_{\mathbb{S}^{d-1}} D^p(\mathcal{R}_{\mathbf{u}}\eta * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} \\ &= G_\sigma \text{SD}_p(\mu, \eta) + G_\sigma \text{SD}_p(\eta, \nu), \end{aligned}$$

where inequality in (\star) follows from the application of Minkowski inequality.

A.2 Proof of Theorem 3.2: $G_\sigma \text{SD}_p$ metrizes the weak topology

The proof is done by double implications and the technical material relies on the continuous mapping theorem (Athreya & Lahiri, 2006) and bounded convergence theorem for the first direct implication “ \Rightarrow ”. The second one, “ \Leftarrow ”, is based on the fact that weak convergence is equivalent to the convergence corresponding to Lévy-Prokhorov distance (Huber, 2011)

“ \Rightarrow ” Assume that $\mu_k \Rightarrow \mu$. Fix $\mathbf{u} \in \mathbb{S}^{d-1}$, the mapping $\mathbf{u} \mapsto \mathcal{R}_{\mathbf{u}}$ is continuous from \mathbb{R}^d to \mathbb{R} , then an application of continuous mapping theorem (Athreya & Lahiri, 2006) entails that $\mathcal{R}_{\mathbf{u}}\mu_k \Rightarrow \mathcal{R}_{\mathbf{u}}\mu$. By Lévy’s continuity theorem (Athreya & Lahiri, 2006) $\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma \Rightarrow \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$. Therefore, $\lim_{k \rightarrow \infty} D(\mathcal{R}_{\mathbf{u}}\mu_k, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) = 0$. Since we suppose that the divergence D is bounded, then there exists $K \geq 0$ such that for any k , $D^p(\mathcal{R}_{\mathbf{u}}\mu_k, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) \leq K$. An application of bounded convergence theorem yields

$$\lim_{k \rightarrow \infty} G_\sigma \text{SD}_p(\mu_k, \mu) = \left(\int_{\mathbb{S}^{d-1}} \lim_{k \rightarrow \infty} D^p(\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} = 0.$$

“ \Leftarrow ” (By contrapositive). Suppose that μ_k doesn’t converge weakly to μ and assume that $\lim_{k \rightarrow \infty} G_\sigma \text{SD}_p(\mu_k, \mu) = 0$. On one hand, since \mathbb{R}^d is a complete separable space then the weak convergence is equivalent to the convergence corresponding to Lévy-Prokhorov distance Λ defined as: The Lévy-Prokhorov distance $\Lambda(\eta, \zeta)$ between $\eta, \zeta \in \mathcal{P}((E, \rho), \mathcal{T})$ (space of probability measures on a measurable metric space) is given by:

$$\Lambda(\eta, \zeta) = \inf_{\varepsilon > 0} \{ \eta(A) < \zeta(A^\varepsilon) + \varepsilon, \quad \zeta(A) < \eta(A^\varepsilon) + \varepsilon, \quad \text{for all } A \in \mathcal{T} \}, \quad \text{where } A^\varepsilon = \{x \in E : \rho(x, A) < \varepsilon\}.$$

Hence there exists $\varepsilon > 0$ and a subsequence $\{\mu_{s(k)}\}_{k \in \mathbb{N}}$ such that $\Lambda(\mu_{s(k)}, \mu) > \varepsilon$. On the other hand, we have $\lim_{k \rightarrow \infty} G_\sigma \text{SD}_p(\mu_{s(k)}, \mu) = 0$, that is equivalent to $\{D(\mathcal{R}_{\mathbf{u}}\mu_{s(k)} * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma)\}_k$ converges to 0 in $L^p(\mathbb{S}^{d-1}) = \{f : \mathbb{S}^{d-1} \rightarrow \mathbb{R} \mid \int_{\mathbb{S}^{d-1}} f(\mathbf{u}) u_d(\mathbf{u}) d\mathbf{u} < \infty\}$. Since the L^p -convergence entails the point-wise convergence (Khoshnevisan, 2007), there exists a subsequence $\{\mu_{s(t(k))}\}_k$ such that $\lim_{k \rightarrow \infty} D(\mathcal{R}_{\mathbf{u}}\mu_{s(t(k))} * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) = 0$ almost everywhere for all $\mathbf{u} \in \mathbb{S}^{d-1}$. Recall that the divergence D metrizes the weak convergence in $\mathcal{P}(\mathbb{R})$ then $\mathcal{R}_{\mathbf{u}}\mu_{s(t(k))} * \mathcal{N}_\sigma \Rightarrow \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$ almost everywhere for all $\mathbf{u} \in \mathbb{S}^{d-1}$. Therefore, $\mathcal{R}_{\mathbf{u}}\mu_{s(t(k))} \Rightarrow \mathcal{R}_{\mathbf{u}}\mu$ almost everywhere for all $\mathbf{u} \in \mathbb{S}^{d-1}$. Using Cramér-Wold device (Huber, 2011), we get $\mu_{s(t(k))} \Rightarrow \mu$. Since the Lévy-Prokhorov distance metrizes the weak convergence, it entails that $\lim_{k \rightarrow \infty} \Lambda(\mu_{s(t(k))}, \mu_k) = 0$, that contradicts the fact that $\Lambda(\mu_{s(k)}, \mu) > \varepsilon$. We then conclude by contrapositive that $\mu_k \Rightarrow \mu$.

A.3 Proof of Proposition 3.3: $G_\sigma \text{SD}_p$ is lower semi-continuous

Recall that the base divergence D is lower semi-continuous w.r.t. the weak topology in $\mathcal{P}(\mathbb{R})$, namely for every sequence of measures $\{\mu'_k\}_{k \in \mathbb{N}}$ and $\{\nu'_k\}_{k \in \mathbb{N}}$ in $\mathcal{P}(\mathbb{R})$ such that $\mu'_k \Rightarrow \mu'$ and $\nu'_k \Rightarrow \nu'$, one has $D(\mu', \nu') \leq \liminf_{k \rightarrow \infty} D(\mu'_k, \nu'_k)$.

Now, let $\{\mu_k\}_{k \in \mathbb{N}}$ and $\{\nu_k\}_{k \in \mathbb{N}}$ are two sequences of measure in $\mathcal{P}_p(\mathbb{R}^d)$ such that $\mu_k \Rightarrow \mu$ and $\nu_k \Rightarrow \nu$.

By continuous mapping theorem (Bowers & Kalton, 2014) and Levy's continuity theorem, we obtain $\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma \Rightarrow \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$ and $\mathcal{R}_{\mathbf{u}}\nu_k * \mathcal{N}_\sigma \Rightarrow \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma$ for all $\mathbf{u} \in \mathbb{S}^{d-1}$. Since the base divergence D is a lower semi-continuous with respect to weak topology in $\mathcal{P}(\mathbb{R})$, then

$$D^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) \leq \left(\liminf_{k \rightarrow \infty} D(\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu_k * \mathcal{N}_\sigma) \right)^p \leq \liminf_{k \rightarrow \infty} D^p(\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu_k * \mathcal{N}_\sigma).$$

It gives

$$G_\sigma \text{SD}_p(\mu, \nu) \leq \left(\int_{\mathbb{S}^{d-1}} \liminf_{k \rightarrow \infty} D^p(\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu_k * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p}.$$

Furthermore, by application of Fatou's lemma (Bowers & Kalton, 2014), we get

$$G_\sigma \text{SD}_p(\mu, \nu) \leq \liminf_{k \rightarrow \infty} \left(\int_{\mathbb{S}^{d-1}} D^p(\mathcal{R}_{\mathbf{u}}\mu_k * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu_k * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} = \liminf_{k \rightarrow \infty} G_\sigma \text{SD}_p(\mu_k, \nu_k),$$

which is the desired result.

A.4 Proofs of statistical properties

A.4.1 Proof of Lemma 3.5: $\mathcal{R}_{\mathbf{u}}\hat{\mu}_n * \mathcal{N}_\sigma$ is an average of Gaussian mixture

Straighforwardly, for every Borelian $I \in \mathcal{B}(\mathbb{R})$, we have

$$\begin{aligned} \mathcal{R}_{\mathbf{u}}\hat{\mu}_n * \mathcal{N}_\sigma(I) &= \int_r \int_s \mathbf{1}_I(r+s) d\left\{ \frac{1}{n} \sum_{i=1}^n \delta_{\mathbf{u}^\top X_i} \right\}(r) d\mathcal{N}_\sigma(s) \\ &= \frac{1}{n} \sum_{i=1}^n \int_s \mathbf{1}_I(\mathbf{u}^\top X_i + s) f_{\mathcal{N}_\sigma}(s) ds \\ &= \frac{1}{n} \sum_{i=1}^n \int_{s'} \mathbf{1}_I(s') f_{\mathcal{N}_\sigma}(s' - \mathbf{u}^\top X_i) ds' \\ &= \frac{1}{n} \sum_{i=1}^n \int_{s'} \mathbf{1}_I(s') f_{\mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)}(s') ds' \quad (\text{since } f_{\mathcal{N}_\sigma}(s' - \mathbf{u}^\top X_i) = f_{\mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)}(s')) \\ &= \frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)(I). \end{aligned}$$

Thanks to Theorem of Cramér and Wold (Cramér & Wold, 1936), we conclude the equality between the measures $\mathcal{R}_{\mathbf{u}}\hat{\mu}_n * \mathcal{N}_\sigma = \frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)$.

A.4.2 Proof of Proposition 3.8

Let us give first the overall structure of the proof. We use frequently the triangle inequality for Wasserstein distances between the quantities $\hat{\mu}_n$, $\frac{1}{n} \mathcal{N}_\sigma(\mathbf{u}^\top X_i, \sigma^2)$ and $\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$. We then obtain two quantities, **I** and **II** (see below for explicit), bounding $\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\mu}_n, \mu)]$. To control **I** bound, we use a well known converging bound in Fournier & Guillin (2015) of Wasserstein distance between empirical and true measure. For **II** bound, we consider maximal TV-coupling in Villani (2009) and use result of the $2p$ -moment of absolute Gaussian random variable founded in Winkelbauer (2014).

On one hand, using triangle inequality of Wasserstein distance, we have

$$\begin{aligned} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\mu}_n, \mu)] &= \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \left[\left(\int_{\mathbb{S}^{d-1}} W_p^p(\hat{\mu}_n, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} \right] \\ &\leq \left(\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \left[\int_{\mathbb{S}^{d-1}} W_p^p(\hat{\mu}_n, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) u_d(\mathbf{u}) d\mathbf{u} \right] \right)^{1/p} \\ &\leq \left(\int_{\mathbb{S}^{d-1}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [W_p^p(\hat{\mu}_n, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma)] u_d(\mathbf{u}) d\mathbf{u} \right)^{1/p} \\ &\leq (\mathbf{I} + \mathbf{II})^{1/p} \end{aligned}$$

where

$$\mathbf{I} \triangleq 2^{p-1} \int_{\mathbb{S}^{d-1}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_{\sigma^n}} \left[\mathbf{W}_p^p \left(\hat{\mu}_n, \frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) \right] u_d(\mathbf{u}) d\mathbf{u}$$

and

$$\mathbf{II} \triangleq 2^{p-1} \int_{\mathbb{S}^{d-1}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_{\sigma^n}} \left[\mathbf{W}_p^p \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2), R_{\mathbf{u}} \mu * \mathcal{N}_\sigma \right) \right] u_d(\mathbf{u}) d\mathbf{u}$$

The proof is based on two steps to control the quantities \mathbf{I} and \mathbf{II} .

Step 1: Control of \mathbf{I} .

Let us state the following lemma:

Lemma A.1 (See proof of Theorem 1 in Fournier & Guillin (2015)). *Let $\eta \in \mathcal{P}(\mathbb{R})$ and let $p \geq 1$. Assume that $M_q(\eta) < \infty$ for some $q > p$. There exists a constant $C_{p,q}$ depending only on p, q such that, for all $n \geq 1$,*

$$\mathbf{E}[\mathbf{W}_p^p(\hat{\eta}_n, \eta)] \leq C_{p,q} M_q(\eta)^{p/q} \Delta_n(p, q),$$

where

$$\Delta_n(p, q) = \begin{cases} n^{-1/2} \mathbf{1}_{q > 2p}, \\ n^{-1/2} \log(n) \mathbf{1}_{q=2p} \\ n^{-(q-p)/q} \mathbf{1}_{p < q < 2p}. \end{cases}.$$

We note that $\hat{\mu}_n$ is an empirical version of the Gaussian mixture $\frac{1}{n} \sum_{i=1}^n \mathcal{N}(u^\top X_i, \sigma^2)$. Then, by application of Lemma A.1, we get

$$\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_{\sigma^n}} \left[\mathbf{W}_p^p \left(\hat{\mu}_n, \frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) \right] \leq C_{p,q} \mathbf{E}_{\mu^{\otimes n}} \left[M_q^{p/q} \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) \right] \Delta_n(p, q).$$

Let us first upper bound the q -th moment of $M_q \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right)$, for all $q \geq 1$. For all $\mathbf{u} \in \mathbb{S}^{d-1}$, we have

$$M_q \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) = \int_{\mathbb{R}} |t|^q d \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) (t) = \frac{1}{n} \sum_{i=1}^n M_q(|Z_{i,\mathbf{u}}|^q),$$

where $Z_{i,\mathbf{u}} \sim \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)$. By Equation (17) in Winkelbauer (2014) we have

$$M_q \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) = \frac{1}{n} \frac{2^{q/2} \sigma^q}{\sqrt{\pi}} \Gamma\left(\frac{q+1}{2}\right) \sum_{i=1}^n {}_1F_1\left(-\frac{q}{2}, \frac{1}{2}; \frac{-(\mathbf{u}^\top X_i)^2}{2\sigma^2}\right).$$

Since X_1, \dots, X_n are i.i.d samples from μ , it yields

$$\begin{aligned} \mathbf{E}_{\mu^{\otimes n}} \left[M_q^{p/q} \left(\frac{1}{n} \sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2) \right) \right] &= \frac{2^{q/2} \sigma^q}{\sqrt{\pi}} \Gamma\left(\frac{q+1}{2}\right) \mathbf{E}_\mu \left[{}_1F_1\left(-\frac{q}{2}, \frac{1}{2}; \frac{-(\mathbf{u}^\top X)^2}{2\sigma^2}\right) \right] \quad (X \sim \mu) \\ &= \frac{2^{q/2} \sigma^q}{\sqrt{\pi}} \Gamma\left(\frac{q+1}{2}\right) \sum_{k=0}^{\infty} \frac{\left(-\frac{q}{2}\right)_k}{\left(\frac{1}{2}\right)_k} \frac{(-1)^k}{(2\sigma^2)^k k!} \mathbf{E}_\mu[(\mathbf{u}^\top X)^{2k}] \\ &\leq \frac{2^{q/2} \sigma^q}{\sqrt{\pi}} \Gamma\left(\frac{q+1}{2}\right) \sum_{k=0}^{\infty} \frac{\left(-\frac{q}{2}\right)_k}{\left(\frac{1}{2}\right)_k} \frac{(-1)^k}{(2\sigma^2)^k k!} M_{2k}(\mu). \end{aligned}$$

Setting $q = 2p$ we have $\Delta_n(p, q) = \frac{\log n}{n}$, then

$$\mathbf{I} \leq 2^{2p-1} C_p \frac{\sigma^{2p}}{\sqrt{\pi}} \Gamma\left(\frac{2p+1}{2}\right) \sum_{k=0}^{\infty} \frac{(-p)_k}{\left(\frac{1}{2}\right)_k} \frac{(-1)^k}{(2\sigma^2)^k k!} M_{2k}(\mu) \frac{\log(n)}{n}.$$

Step 2: Control of II.

We follow the lines of proofs of Proposition 1 in Goldfeld et al. (2020) and Theorem 2 in Nietert et al. (2021). Using a coupling $\hat{\mu}_n$ and $\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$ via the maximal TV-coupling (see Theorem 6.15 in Villani (2009)), the control of the total variation of the Wasserstein distance, we get for any fixed $\mathbf{u} \in \mathbb{S}^{d-1}$

$$\mathbb{W}_p^p\left(\frac{1}{n}\sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2), \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma\right) \leq 2^{p-1} \int_{\mathbb{R}} |t|^p |h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t)| dt,$$

where $h_{n,\mathbf{u}}$ and $g_{\mathbf{u}}$ are the densities associated with μ_n and $\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma$, respectively. Let $f_{\sigma,\vartheta}$ the probability density function of $\mathcal{N}_{\sigma,\vartheta}$, i.e, $f_{\sigma,\vartheta}(t) = \frac{1}{\sqrt{2\pi(\sigma\vartheta)^2}} e^{-\frac{t^2}{2(\sigma\vartheta)^2}}$ for $\vartheta > 0$ to be specified later. An application of Cauchy-Schwarz inequality gives

$$\begin{aligned} \mathbf{E}_{\mu^{\otimes n}|\mathcal{N}_\sigma^{\otimes n}} \left[\mathbb{W}_p^p\left(\frac{1}{n}\sum_{i=1}^n \mathcal{N}(\mathbf{u}^\top X_i, \sigma^2), \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma\right) \right] \\ \leq 2^{p-1} \mathbf{E}_{\mu^{\otimes n}|\mathcal{N}_\sigma^{\otimes n}} \int_{\mathbb{R}} |t|^p \sqrt{f_{\sigma,\vartheta}(t)} \frac{|h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t)|}{\sqrt{f_{\sigma,\vartheta}(t)}} dt \\ \leq 2^{p-1} \mathbf{E}_{\mu^{\otimes n}|\mathcal{N}_\sigma^{\otimes n}} \left(\int_{\mathbb{R}} |t|^{2p} f_{\sigma,\vartheta}(t) dt \right)^{\frac{1}{2}} \left(\int_{\mathbb{R}} \frac{(h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t))^2}{f_{\sigma,\vartheta}(t)} dt \right)^{\frac{1}{2}} \\ \leq 2^{p-1} \left(\int_{\mathbb{R}} |t|^{2p} f_{\sigma,\vartheta}(t) dt \right)^{\frac{1}{2}} \left(\int_{\mathbb{R}} \mathbf{E}_{\mu^{\otimes n}|\mathcal{N}_\sigma^{\otimes n}} \frac{(h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t))^2}{f_{\sigma,\vartheta}(t)} dt \right)^{\frac{1}{2}}. \end{aligned}$$

Note that $\int_{\mathbb{R}} |t|^{2p} f_{\sigma,\vartheta}(t) dt$ is the $2p$ -th moment of $|\mathcal{N}_{\sigma,\vartheta}(t)|$ equals to (see Equation (18) in Winkelbauer (2014))

$$\int_{\mathbb{R}} |t|^{2p} f_{\sigma,\vartheta}(t) dt = \frac{(\sigma\vartheta)^{2p} 2^p}{\sqrt{\pi}} \Gamma\left(\frac{2p+1}{2}\right).$$

Moreover,

$$h_{n,\mathbf{u}}(t) = \frac{1}{n} \sum_{i=1}^n d\mathcal{N}(\mathbf{u}^\top X_i, \sigma^2)(t) = \frac{1}{n} \sum_{i=1}^n f_{\sigma,\vartheta}(t - \mathbf{u}^\top X_i),$$

It is clear to see that $h_{n,\mathbf{u}}(t)$ is a sum of i.i.d. terms with expectation $g_{\mathbf{u}}(t)$, which implies

$$\begin{aligned} \mathbf{E}_{\mu^{\otimes n}|\mathcal{N}_\sigma^{\otimes n}} [(h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t))^2] &= \mathbf{V}_{\mu^{\otimes n}} \left[\frac{1}{n} \sum_{i=1}^n f_{\sigma,\vartheta}(t - \mathbf{u}^\top X_i) \right] \\ &= \frac{1}{n} \mathbf{V}_{\mu} [f_{\sigma,\vartheta}(t - \mathbf{u}^\top X)] \\ &\leq \frac{1}{n} \mathbf{E}_{\mu} [(f_{\sigma,\vartheta}(t - \mathbf{u}^\top X))^2] \\ &\leq \frac{(2\pi\sigma^2)^{-1}}{n} \mathbf{E}_{\mu} [e^{\frac{-1}{\sigma^2}(t - \mathbf{u}^\top X)^2}]. \end{aligned}$$

Now

$$\mathbf{E}_{\mu} [e^{\frac{-(t - \mathbf{u}^\top X)^2}{\sigma^2}}] = \int_{\|x\| \leq \frac{|t|}{2}} e^{\frac{-1}{\sigma^2}(t - \mathbf{u}^\top x)^2} d\mu(x) + \int_{\|x\| > \frac{|t|}{2}} e^{\frac{-1}{\sigma^2}(t - \mathbf{u}^\top x)^2} d\mu(x).$$

Remark that when $\|x\| \leq \frac{|t|}{2}$, then $(t - \mathbf{u}^\top X)^2 \geq |t|^2 - |\mathbf{u}^\top x|^2 \geq |t|^2 - \|x\|^2$ (since $\|u\|^2 = 1$). We get $(t - \mathbf{u}^\top X)^2 \geq \frac{|t|^2}{4}$ and hence

$$\int_{\|x\| \leq \frac{|t|}{2}} e^{\frac{-1}{\sigma^2}(t - \mathbf{u}^\top x)^2} d\mu(x) \leq e^{\frac{-t^2}{4\sigma^2}} \text{ and } \int_{\|x\| > \frac{|t|}{2}} e^{\frac{-1}{\sigma^2}(t - \mathbf{u}^\top x)^2} d\mu(x) \leq \mathbf{P}[\|X\| > \frac{|t|}{2}]$$

This gives,

$$\int_{\mathbb{R}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \frac{(h_{n,\mathbf{u}}(t) - g_{\mathbf{u}}(t))^2}{f_{\sigma,\vartheta}(t)} dt \leq \frac{(2\pi\sigma^2)^{-1}(\sqrt{2\pi}\sigma\vartheta)}{n} \left(\int_{\mathbb{R}} e^{\frac{t^2}{2(\sigma\vartheta)^2}} e^{\frac{-t^2}{4\sigma^2}} dt + \int_{\mathbb{R}} e^{\frac{t^2}{2(\sigma\vartheta)^2}} \mathbf{P}[\|X\| > \frac{|t|}{2}] dt \right).$$

Note that the integral $\int_{\mathbb{R}} e^{\frac{t^2}{2(\sigma\vartheta)^2}} e^{\frac{-t^2}{4\sigma^2}} dt = \int_{\mathbb{R}} e^{-\left(\frac{1}{2} - \frac{1}{\vartheta^2}\right) \frac{t^2}{2\sigma^2}} dt$ is finite if and only if $\frac{1}{2} - \frac{1}{\vartheta^2} > 0$ namely $\vartheta > \sqrt{2}$ and its value is given by

$$\int_{\mathbb{R}} e^{\frac{t^2}{2(\sigma\vartheta)^2}} e^{\frac{-t^2}{4\sigma^2}} dt = \sqrt{\frac{2\pi\sigma^2}{\frac{1}{2} - \frac{1}{\vartheta^2}}} = \sqrt{\frac{4\pi\sigma^2\vartheta^2}{\vartheta^2 - 2}}.$$

For the second integral

$$\int_{\mathbb{R}} e^{\frac{t^2}{2(\sigma\vartheta)^2}} \mathbf{P}[\|X\| > \frac{|t|}{2}] dt = 2 \int_0^\infty e^{\frac{t^2}{2(\sigma\vartheta)^2}} \mathbf{P}[\|X\| > \frac{t}{2}] dt = 4 \int_0^\infty e^{\frac{2\xi^2}{\sigma^2\vartheta^2}} \mathbf{P}[\|X\| > \xi] d\xi$$

Then,

$$\mathbf{II} \leq n^{-1/2} 4^{p-1} \left\{ (2\pi\sigma^2)^{-1} (\sqrt{2\pi}\sigma\vartheta) \frac{(\sigma\vartheta)^{2p} 2^p}{\sqrt{\pi}} \Gamma\left(\frac{2p+1}{2}\right) \right\}^{\frac{1}{2}} \left(\sqrt{\frac{4\pi\sigma^2\vartheta^2}{\vartheta^2 - 2}} + 4 \int_0^\infty e^{\frac{2\xi^2}{\sigma^2\vartheta^2}} \mathbf{P}[\|X\| > \xi] d\xi \right)^{\frac{1}{2}}.$$

this gives the desired result using the fact that $(a+b)^{1/p} \leq a^{1/p} + b^{1/p}$, for $a, b \geq 0$.

A.4.3 Proof of Proposition 3.11

Using triangle inequality, we have

$$W_p(\hat{\mu}_n, \hat{\nu}_n) \leq W_p(\hat{\mu}_n, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) + W_p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) + W_p(\mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma, \hat{\nu}_n).$$

and then

$$W_p^p(\hat{\mu}_n, \hat{\nu}_n) \leq 3^{p-1} \{ W_p^p(\hat{\mu}_n, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma) + W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) + W_p^p(\mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma, \hat{\nu}_n) \}.$$

This implies that

$$\begin{aligned} & \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\mu}_n, \hat{\nu}_n)] \\ & \leq 3^{1-\frac{1}{p}} G_\sigma \text{SW}_p(\mu, \nu) + 3^{1-\frac{1}{p}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\mu}_n, \mu)] + 3^{1-\frac{1}{p}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\nu}_n, \nu)]. \end{aligned}$$

By application of Proposition 3.8, it yields This gives that

$$\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{G}_\sigma \text{SW}_p(\hat{\mu}_n, \hat{\nu}_n)] \leq 3^{1-\frac{1}{p}} G_\sigma \text{SW}_p(\mu, \nu) + 3\Xi_{p,\sigma,\vartheta} \frac{1}{n^{1/2p}} + 3^{1-\frac{1}{p}} (\Upsilon_{p,\sigma,\mu} + \Upsilon_{p,\sigma,\nu}) \frac{(\log n)^{1/p}}{n^{1/p}}$$

This ends the proof of the first statement in Proposition 3.11. For the second one, we also use a triangle inequality

$$W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) \leq 3^{p-1} \{ W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_\sigma, \hat{\mu}_n) + W_p^p(\hat{\mu}_n, \hat{\nu}_n) + W_p^p(\hat{\nu}_n, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_\sigma) \}.$$

Then we control each term as we did before.

A.5 Proof of Proposition 3.12: projection complexity

Using Holder's inequality, we have

$$\begin{aligned} \mathbf{E}_{u_d^{\otimes L}} [|\widehat{G_\sigma \text{SD}}_p^p(\mu, \nu) - G_\sigma \text{SD}_p^p(\mu, \nu)|] & \leq \left(\mathbf{E}_{u_d^{\otimes L}} [|\widehat{G_\sigma \text{SD}}_p^p(\mu, \nu) - G_\sigma \text{SD}_p^p(\mu, \nu)|^2] \right)^{1/2} \\ & = \left(\mathbf{V}_{u_d^{\otimes L}} [|\widehat{G_\sigma \text{SD}}_p^p(\mu, \nu)|] \right)^{1/2} \\ & = \frac{A(p, \sigma)}{L^{1/2}}. \end{aligned}$$

A.6 Proof of Corollary 3.13: overall complexity ($p = 1$)

By application of triangle inequality, one has

$$|\widehat{\hat{G}}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - G_\sigma \text{SW}(\mu, \nu)| \leq |\widehat{\hat{G}}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - \hat{G}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n)| + |\hat{G}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - G_\sigma \text{SW}(\mu, \nu)|$$

Using Proposition 3.12, we have

$$\mathbf{E}_{u_d^{\otimes L}} [|\widehat{\hat{G}}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - \hat{G}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n)|] \leq \frac{\hat{A}_\sigma}{\sqrt{L}} := \frac{\{\mathbf{V}_{\mathbf{u} \sim u_d}[\mathbf{W}(\hat{\mu}_n, \hat{\nu}_n)]\}^{1/2}}{\sqrt{L}}.$$

Using Proposition 3.11 for $p = 1$ we get,

$$\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [|\hat{G}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - G_\sigma \text{SW}(\mu, \nu)|] \leq 3\Xi_{1, \sigma, \vartheta} \frac{1}{\sqrt{n}} + (\Upsilon_{1, \sigma, \mu} + \Upsilon_{1, \sigma, \nu}) \frac{\log n}{n}.$$

Therefore, by applying the expectations with respect to the projection and sampling we obtain

$$\begin{aligned} \mathbf{E}_{u_d^{\otimes L}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [|\widehat{\hat{G}}_\sigma \text{SW}(\hat{\mu}_n, \hat{\nu}_n) - G_\sigma \text{SW}(\mu, \nu)|] \\ \leq \frac{1}{\sqrt{L}} \mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{A}_\sigma] + 3\Xi_{1, \sigma, \vartheta} \frac{1}{\sqrt{n}} + (\Upsilon_{1, \sigma, \mu} + \Upsilon_{1, \sigma, \nu}) \frac{\log n}{n}. \end{aligned}$$

By Jensen inequality, we have

$$\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\hat{A}_\sigma] \leq \{\mathbf{E}_{\mu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} \mathbf{E}_{\nu^{\otimes n} | \mathcal{N}_\sigma^{\otimes n}} [\mathbf{V}_{\mathbf{u} \sim u_d}[\mathbf{W}(\hat{\mu}_n, \hat{\nu}_n)]]\}^{1/2}.$$

A.7 Proof of Proposition 3.14

For all $\mathbf{u} \in \mathbb{S}^{d-1}$ we have $\mathcal{R}_{\mathbf{u}}\mu, \mathcal{R}_{\mathbf{u}}\nu \in \mathcal{P}(\mathbb{R})$. By application of the inequality of noise level satisfied by D in one dimension we get

$$D^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_2}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_2}) \leq D^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_1}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_1}).$$

Then, computing the expectation over the projections \mathbf{u} since the divergence is non-negative concludes the proof.

A.8 Proof of Proposition 3.16: relation between $G_\sigma \text{SW}^p(\mu, \nu)$ under two noise levels

First, using the contractive property of convolution (see Lemma 3 in Nietert et al. (2021)), stating that for any probability measure $\alpha \in \mathcal{P}(\mathbb{R})$, $W_p(\mu * \alpha, \nu * \alpha) \leq W_p(\mu, \nu)$. Hence $W_p^p(\mu * \mathcal{N}_{\sigma_2}, \nu * \mathcal{N}_{\sigma_2}) \leq W_p^p(\mu * \mathcal{N}_{\sigma_1}, \nu * \mathcal{N}_{\sigma_1})$. Now using Proposition 3.14 of the order relation satisfied by $G_\sigma \text{SW}^p$ yields

$$G_{\sigma_2} \text{SW}_p(\mu, \nu) \leq G_{\sigma_1} \text{SW}_p(\mu, \nu).$$

In the other direction, we have that $\mathcal{N}_{\sigma_2} = \mathcal{N}_{\sigma_1} * \mathcal{N}_{\sqrt{\sigma_2^2 - \sigma_1^2}}$ (similarly for \mathcal{N}_{σ_1}). Setting the following random variables: $X_{\mathbf{u}} \sim \mathcal{R}_{\mathbf{u}}\mu, Y_{\mathbf{u}} \sim \mathcal{R}_{\mathbf{u}}\nu, Z_X \sim \mathcal{N}_{\sigma_1}, Z_Y \sim \mathcal{N}_{\sigma_1}, Z'_X \sim \mathcal{N}_{\sqrt{\sigma_2^2 - \sigma_1^2}}, Z'_Y \sim \mathcal{N}_{\sqrt{\sigma_2^2 - \sigma_1^2}}$. The sliced Wasserstein distance $W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_2}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_2})$ is given as a minimization over couplings $(X_{\mathbf{u}}, Z_X, Z'_X)$ and $(Y_{\mathbf{u}}, Z_Y, Z'_Y)$, namely

$$W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_2}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_2}) = \inf_{\substack{X_{\mathbf{u}}, Z_X, Z'_X \\ Y_{\mathbf{u}}, Z_Y, Z'_Y}} \mathbf{E}[|((X_{\mathbf{u}} + Z_X) - (Y_{\mathbf{u}} + Z_Y)) + (Z'_X - Z'_Y)|^p]$$

Using the inequality $\mathbf{E}[|U + V|^p] - 2^{p-1} \mathbf{E}[|W|^p] \leq 2^{p-1} \mathbf{E}[|U + V + W|^p]$ for any random variables $U, V, W \in \mathbb{L}_p$ integrable, we obtain,

$$2^{p-1} \mathbf{E}[|(X_{\mathbf{u}} + Z_X) - (Y_{\mathbf{u}} + Z_Y) + (Z'_X + Z'_Y)|^p] \geq \mathbf{E}[|(X_{\mathbf{u}} + Z_X) - (Y_{\mathbf{u}} + Z_Y)|^p] - 2^{p-1} \mathbf{E}[|(Z'_X - Z'_Y)|^p].$$

Hence,

$$\begin{aligned} 2^{p-1}W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_2}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_2}) &\geq \inf \left(\mathbf{E}[|(X_{\mathbf{u}} + Z_X) - (Y_{\mathbf{u}} + Z_Y)|^p] - 2^{p-1}\mathbf{E}[|(Z'_X - Z'_Y)|^p] \right) \\ &\geq W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_1}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_1}) - 2^{p-1} \sup \mathbf{E}[|(Z'_X - Z'_Y)|^p] \\ &\geq W_p^p(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma_1}, \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma_1}) - 2^{2p} \sup \mathbf{E}[|(Z'_X)|^p]. \end{aligned}$$

Hence,

$$G_{\sigma_1}SW_p(\mu, \nu) \leq 2^{1-\frac{1}{p}} G_{\sigma_2}SW_p(\mu, \nu) + 4(\sup \mathbf{E}[|(Z'_X)|^p])^{1/p}.$$

Finally, for any $p \geq 1$ the p -th moment of $|\mathcal{N}_{\sigma}|^p$ satisfies $\mathbf{E}[|\mathcal{N}_{\sigma}|^p] = \frac{2^p \Gamma((p+1)/2)}{\Gamma(1/2)} \sigma^{2p} \leq 2^{p/2} \sigma^{2p}$, then

$$G_{\sigma_1}SW_p(\mu, \nu) \leq 2^{1-\frac{1}{p}} G_{\sigma_2}SW_p(\mu, \nu) + 2^{\frac{5}{2}}(\sigma_2^2 - \sigma_1^2),$$

and concludes the proof.

A.9 Proof of Proposition 3.17: continuity of the smoothed Gaussian sliced Wasserstein w.r.t. σ

From Lemma 1 in (Nietert et al., 2021), we know that the Gaussian-smoothed Wasserstein is continuous with respect to σ , for any distribution $\mathcal{R}_{\mathbf{u}}\nu$ and $\mathcal{R}_{\mathbf{u}}\mu$. In addition, for any \mathbf{u} , we have $W_p(\mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma}, \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}) \leq W_p(\mathcal{R}_{\mathbf{u}}\nu, \mathcal{R}_{\mathbf{u}}\mu)$. Then by applying Lebesgue's dominated convergence theorem (Bowers & Kalton, 2014) to the above inequality with $W_p(\mathcal{R}_{\mathbf{u}}\nu, \mathcal{R}_{\mathbf{u}}\mu)$ as a dominating function, that is u_d -almost everywhere integrable because both measures are in $\mathcal{P}_p(\mathbb{R}^d)$, we then conclude that the Gaussian-smoothed SWD is continuous w.r.t. σ .

A.10 Proof of Proposition 3.18: continuity of the smoothed sliced squared-MMD w.r.t. σ

Let us first recall the definition of the MMD divergence. Let $k : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$ be a measurable bounded kernel on \mathbb{R} and consider the reproducing kernel Hilbert space (RKHS) \mathcal{H}_k associated with k and equipped with inner product $\langle \cdot, \cdot \rangle_{\mathcal{H}_k}$ and norm $\|\cdot\|_{\mathcal{H}_k}$. Let $\mathcal{P}_{\mathcal{H}_k}(\mathbb{R})$ be the set of probability measures η such that $\int_{\mathbb{R}} \sqrt{k(t, t)} d\eta(x) < \infty$. The kernel mean embedding is defined as $\Phi_k(\eta) = \int_{\mathbb{R}} k(\cdot, t) d\eta(t)$. The squared-maximum mean discrepancy between $\eta, \zeta \in \mathcal{P}(\mathbb{R})$ denoted as $\text{MMD} : \mathcal{P}_{\mathcal{H}_k}(\mathbb{R}) \times \mathcal{P}_{\mathcal{H}_k}(\mathbb{R}) \rightarrow \mathbb{R}_+$ is expressed as the distance between two such kernel mean embeddings. It is defined as Gretton et al. (2012)

$$\text{MMD}^2(\eta, \zeta) = \|\Phi_k(\eta) - \Phi_k(\zeta)\|_{\mathcal{H}_k}^2 = \mathbf{E}_{T, T' \sim \eta}[k(T, T')] - 2\mathbf{E}_{T \sim \eta, R \sim \zeta}[k(T, R)] + \mathbf{E}_{R, R' \sim \zeta}[k(R, R')]$$

where T and T' are independent random variables drawn according to η , R and R' are independent random variables drawn according to ζ , and T is independent of R . We define the Gaussian Smoothed Sliced squared-MMD as follows:

$$\begin{aligned} G_{\sigma} \text{MMD}^2(\mu, \nu) &= \int_{\mathbb{S}^{d-1}} \|\Phi_k(\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}) - \Phi_k(\mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma})\|_{\mathcal{H}_k}^2 u_d(\mathbf{u}) d\mathbf{u} \\ &= \int_{\mathbb{S}^{d-1}} \left(\mathbf{E}_{T, T' \sim \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}}[k(T, T')] - 2\mathbf{E}_{T \sim \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}, R \sim \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma}}[k(T, R)] \right. \\ &\quad \left. + \mathbf{E}_{R, R' \sim \mathcal{R}_{\mathbf{u}}\nu * \mathcal{N}_{\sigma}}[k(R, R')] \right) u_d(\mathbf{u}) d\mathbf{u}. \end{aligned}$$

From the definition of the smoothed sliced squared-MMD, we have

$$\begin{aligned} \mathbf{E}_{T, T' \sim \mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}}[k(T, T')] &= \iint_{\mathbb{R} \times \mathbb{R}} k(t, t') d\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}(t) d\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}(t') \\ &= \iint_{\mathbb{R} \times \mathbb{R}} \left(\int_{\mathbb{R}} k(t + z, t') d\mathcal{R}_{\mathbf{u}}\mu(z) \mathcal{N}_{\sigma}(t) \right) d\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}(t') \\ &= \iint_{\mathbb{R} \times \mathbb{R}} \left(\int_{\mathbb{R}^d} k(t + \mathbf{u}^{\top} x, t') d\mu(x) \mathcal{N}_{\sigma}(t) \right) d\mathcal{R}_{\mathbf{u}}\mu * \mathcal{N}_{\sigma}(t') \\ &= \iint_{\mathbb{R} \times \mathbb{R}} \iint_{\mathbb{R}^d \times \mathbb{R}^d} k(t + \mathbf{u}^{\top} x, t' + \mathbf{u}^{\top} x') d\mu(x) d\mu(x') d\mathcal{N}_{\sigma}(t) d\mathcal{N}_{\sigma}(t'). \end{aligned}$$

Similarly,

$$\mathbf{E}_{R, R' \sim \mathcal{R}_{\mathbf{u}} \nu * \mathcal{N}_\sigma} [k(R, R')] = \iint_{\mathbb{R} \times \mathbb{R}} \iint_{\mathbb{R}^d \times \mathbb{R}^d} k(r + \mathbf{u}^\top y, r' + \mathbf{u}^\top y') d\nu(y) d\nu(y') d\mathcal{N}_\sigma(r) d\mathcal{N}_\sigma(r')$$

and

$$\mathbf{E}_{T \sim \mathcal{R}_{\mathbf{u}} \mu * \mathcal{N}_\sigma, R \sim \mathcal{R}_{\mathbf{u}} \nu * \mathcal{N}_\sigma} [k(T, R)] = \iint_{\mathbb{R} \times \mathbb{R}} \iint_{\mathbb{R}^d \times \mathbb{R}^d} k(t + \mathbf{u}^\top x, r + \mathbf{u}^\top y) d\mu(x) d\nu(y) d\mathcal{N}_\sigma(t) d\mathcal{N}_\sigma(r).$$

Together the assumption of boundness of the kernel function k and the continuity of integrals, the three latter terms are continuous functions w.r.t. $\sigma \in (0, \infty)$. Again by the boundness of the kernel function k , there exists a positive finite constant C_k such that

$$|\mathbf{E}_{T, T' \sim \mathcal{R}_{\mathbf{u}} \mu * \mathcal{N}_\sigma} [k(T, T')] - 2\mathbf{E}_{T \sim \mathcal{R}_{\mathbf{u}} \mu * \mathcal{N}_\sigma, R \sim \mathcal{R}_{\mathbf{u}} \nu * \mathcal{N}_\sigma} [k(T, R)] + \mathbf{E}_{R, R' \sim \mathcal{R}_{\mathbf{u}} \nu * \mathcal{N}_\sigma} [k(R, R')]| \leq 4C_k.$$

We conclude the continuity of $\sigma \mapsto G_\sigma \text{MMD}^2(\mu, \nu)$ by an application of the continuity of integrals.

B Additional experiments

B.1 Sample complexity on CIFAR dataset

We have also evaluated the sample complexity for the CIFAR dataset by sampling sets of increasing size. Results reported in Figure 1 confirms the findings obtained from the toy dataset.

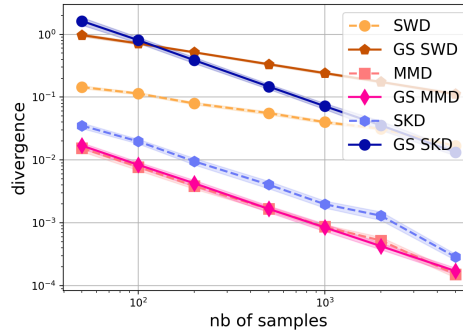


Figure 1: Measuring the divergence between two sets of samples drawn iid from the CIFAR10 dataset. We compare three sliced divergences and their Gaussian smoothed versions with a $\sigma = 3$.

B.2 Identity of indiscernibles

The second experiment aims at checking whether our divergences converge towards a small value when the distributions to be compared are the same. For this, we consider samples from distributions μ and ν chosen as normal distributions with respectively mean $2 \times \mathbf{1}_d$ and $s\mathbf{1}_d$ with varying s (noted as the displacement). Results are depicted in Figure 2. We can see that all methods are able to attain their minimum when $s = 2$. Interestingly, the gap between the Gaussian smoothed and non-smoothed divergences for Wasserstein and Sinkhorn is almost indiscernible as the distance between distribution increases.

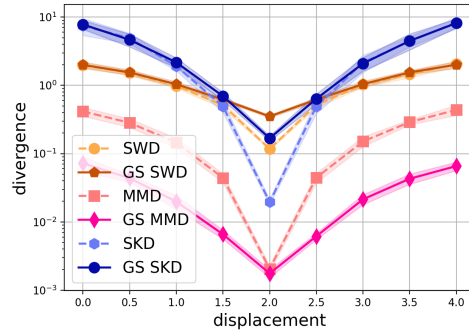


Figure 2: Measuring the divergence between two sets of samples in \mathbb{R}^{50} , one with mean $2\mathbf{1}_d$ and the other with mean $s\mathbf{1}_d$ with increasing s . We compare three sliced divergences and their Gaussian smoothed version with a $\sigma = 3$.