Non-Stationary Bandits under Recharging Payoffs: Improved Planning with Sublinear Regret

Abstract

The stochastic multi-armed bandit setting has been recently studied in the non-stationary regime, where the mean payoff of each action is a non-decreasing function of the number of rounds passed since it was last played. This model captures natural behavioral aspects of the users which crucially determine the performance of recommendation platforms, ad placement systems, and more. Even assuming prior knowledge of the mean payoff functions, computing an optimal planning in the above model is NP-hard, while the state-of-the-art is a $1/4$-approximation algorithm for the case where at most one arm can be played per round. We first focus on the setting where the mean payoff functions are known. In this setting, we significantly improve the best-known guarantees for the planning problem by developing a polynomial-time $(1 - 1/e)$-approximation algorithm (asymptotically and in expectation), based on a novel combination of randomized LP rounding and a time-correlated (interleaved) scheduling method. Furthermore, our algorithm achieves improved guarantees – compared to prior work – for the case where more than one arms can be played at each round. Moving to the bandit setting, when the mean payoff functions are initially unknown, we show how our algorithm can be transformed into a bandit algorithm with sublinear regret.

1 Introduction

In the last two decades, the predominant rise of the social media industry has made the notion of a *newsfeed* an integral part of our lives. In a newsfeed, a user observes a structured sequence of content items (posts, photos etc.) particularly selected by the platform according to the her/his preferences. Apart from social media, an analogous idea – potentially relabeled – also appears in different domains as, for example, "frequently bought together" in e-commerce, "shuffling similar songs" in music recommendation, or "recommended articles" in scholarly literature indexing databases. Whether it is measured in terms of click-rate or time devoted, the high-level objective of newsfeeds is fairly well-known: to maximize the user's engagement with the platform. In many applications, however, achieving this objective is not as simple as identifying the user's "favorite" content, given that her/his satisfaction can depend on the time passed since the same (or similar) content has been observed. As an example, a user's engagement can worsen if a social media feed (resp., a music recommendation platform) constantly presents content from the same source (resp., same artist).

Motivated by such scenarios, researchers have recently studied online decision making problems capturing the notion of "recovering" payoffs, namely, scenarios where the payoff of an action drops to zero after each play and then slowly increases back to a baseline. In the context of online learning, these non-stationary models interpolate between multi-armed bandits, where the environment is assumed to be intact, and reinforcement learning, since the actions may now alter the future environment in a structured manner. These models are wide enough to capture many real-life

applications, yet special enough to accept efficiently computable (near-)optimal solutions. In this direction, the following general model was first introduced (under slightly different assumptions) by Immorlica and Kleinberg [IK18] and has been recently studied by Simchi-Levi et al. [SLZZ21]:

Problem (Multi-armed bandits under recharging payoffs). *We consider a set of* actions *(or* arms*), where each arm i is associated with a* mean payoff function p_i. *For each τ, $p_i(\tau)$ is the expected payoff collected for playing action i, when i has been played before exactly τ rounds (we refer to τ as the* delay *of an arm at a specific round). For each arm, we assume that its payoff function: (a) is monotone non-decreasing in τ, and (b) has a known finite* recovery time*, namely, a specific delay after which the function remains constant. At each round, the player pulls a subset of at most k actions, observes the realized payoffs of the arms played (each computed with respect to the delay of the corresponding action), and collects their sum. The objective is to maximize the expected cumulative payoff collected within an unknown time horizon.*

We denote by k-RB an instance of the problem where at most k arms can be played at each round. Further, we distinguish between the *planning* setting, where the payoff functions are known to the player a priori (thus, payoffs can be considered deterministic), and the *learning* setting, where these are initially unknown and the user assumes semi-bandit feedback on the payoffs of the played actions. In the former case, the goal is to design an efficient algorithm which closely-approximates the optimal planning (since this is generally an NP-hard problem), while in the latter, the objective is to construct a bandit algorithm of sublinear regret, defined as the difference in expected payoff between a planning algorithm and its bandit variant due to the initial absence of information of the latter.

1.1 Background and related work

Immorlica and Kleinberg [IK18] first study the 1-RB problem and provide a $(1 - \epsilon)$-approximation (asymptotically) for the planning setting, under the additional assumption that the payoff functions are weakly concave over the whole time horizon. On the other extreme, Basu et al. [BSSS19] provide a long-run $(1 - 1/e)$-approximation for the case where, after each play, an arm *cannot* be played again for a fixed number of rounds – a problem which can be cast as a special instance of 1-RB using Heaviside step payoff functions.[1] Shortly after the introduction of the problem, a number of special cases and variations have been studied [CCB20, PC21, PBG19, BPCS21, BCMTT20, CCB20, LCCBGB22, APB+21] (we address the reader to Appendix A for an overview).

In their recent work, Simchi-Levi et al. [SLZZ21] study the k-RB problem and prove the first $\mathcal{O}(1)$-approximation guarantee under no additional assumptions. For large k, this result is the state-of-the-art. However, for small k, understanding the magnitude of the constant becomes the primary theoretical question. Specifically, the motivating case of newsfeeds (when, for instance, the content items are presented to a user sequentially) can be modeled as an instance of 1-RB. In this case, the approximation guarantee of [SLZZ21] becomes $1/4$, which is significantly weaker compared to both the $(1 - \epsilon)$-approximation of [IK18] for concave functions and the $(1 - 1/e)$-approximation of [BSSS19] for the most "extreme" example of monotone convex functions, i.e., that of Heaviside step functions. The above observation indicates that either the approximability status of the problem is not well-understood, or that the problem does not gradually become "easier" by increasing the concavity of the payoff functions.

1.2 Technical contributions

In our work, we resolve this discrepancy by designing a long-run $(1 - 1/e)$-approximation algorithm (in expectation) for 1-RB, which improves the state-of-the-art due to [SLZZ21]. Simultaneously, our algorithm enjoys the same asymptotic guarantee of $(1 - \mathcal{O}(1/\sqrt{k}))$ for the general case of k-RB as in Simchi-Levi et al. [SLZZ21] with improved and explicit constants, as opposed to the guarantees of [SLZZ21] which do not come in a closed-form. Our algorithm is based on a novel combination of linear programming (LP) rounding and a time-correlated (interleaved) scheduling method, and is significantly simpler to implement compared to prior work. For the case where the mean payoff functions are initially unknown, we show how our algorithm can be transformed into a bandit algorithm with sublinear regret guarantees.

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[1]Recall that the Heaviside step function is defined such that $H(\tau) = 1$ for $\tau \geq 0$, and $H(\tau) = 0$, otherwise.

1.3 Key ideas and intuition

Henceforth, we refer to any instance where the payoff function of each action i has the form $q_i(\tau) = p_i \cdot H(\tau - d_i)$, where $p_i \in [0, 1]$, $d_i \in \mathbb{N}$, and $H(\cdot)$ the Heaviside step function, as the *Heaviside k-RB* problem. The fact that our $(1 - 1/e)$-approximation for 1-RB matches the state-of-the-art bound obtained for Heaviside 1-RB [BSSS19, PC21] is no coincidence. Our solution technique shows that each k-RB problem is in fact (approximately) hiding a Heaviside k-RB problem.

Structural characterization of a natural LP relaxation. A key idea in both [IK18] and [SLZZ21] is to construct a concave relaxation of the optimal solution. Instead, we take a more direct approach and construct a natural LP relaxation (see Section 3) of the optimal average payoff (our construction uses a uniform upper bound on the recovery time of any arm). By carefully analyzing the structure of our LP, we prove that its extreme point solutions follow a particularly interesting *sparsity pattern*: there exist unique delays $\{\tau_i\}$ associated with the arms (which we refer to as "critical" delays), such that playing each arm i exactly once every τ_i rounds matches the average payoff of the relaxation. As we show, an exception to the above rule can be at most a single arm, which we refer to as *irregular*. The above observation already hints that our problem could potentially be reduced to an instance of Heaviside k-RB for which better approximation guarantees (compared to [SLZZ21]) are known to exist (at least for small k) [BSSS19, PC21].

Improved approximation guarantees for the planning k-RB. Constructing a planning where each arm is played at a rate indicated by its critical delay (as described above) is generally infeasible, since these rates stem from a relaxation of the problem. However, targeting these rates becomes the starting point of our approximation. In Section 4, we design a novel algorithm for the k-RB problem, called "Randomize-Then-Interleave", which starts from computing an optimal solution to our LP relaxation. By combining the sparse structure of this solution with a *randomized rounding* step for determining a critical delay for the (possible) irregular arm, our algorithm produces a "proxy" instance of the Heaviside k-RB problem. Finally, the algorithm applies the *interleaved scheduling* method of [PC21] on the resulting Heaviside k-RB instance. As we prove, the above algorithm achieves a long-run $(1 - \mathcal{O}(1/\sqrt{k}))$-approximation guarantee (in expectation), but with improved and explicit constants compared to [SLZZ21]. Importantly, for the particular case of k-RB, our guarantee is stronger than the one provided in the original work of [PC21]. As a critical step for achieving this improvement, we leverage existing results on the *correlation gap* of the weighted rank function of uniform matroids due to [Yan11] (see Section 2).

For the interesting case of 1-RB, the guarantee of our algorithm becomes 0.316, which already improves on the $1/4$-approximation due to Simchi-Levi et al. [SLZZ21]. However, it is still far from the best known $(1 - 1/e)$-approximation for the special case of Heaviside 1-RB. For this special case, we provide an alternative and tighter analysis of our algorithm, which shows that, indeed, the latter achieves a long-run $(1 - 1/e)$-approximation (in expectation) for 1-RB. Our strategy behind this alternative proof relies on the novel idea of considering parallel (fictitious) copies of the irregular arm, one for each possible critical delay. While the marginal probabilities of each copy being played under the corresponding delay at any round are consistent with the initial LP solution, these events are *mutually exclusive*. For the case of 1-RB, we show that expected payoff collected can only decrease by assuming that the parallel copies were instead independent. Having established that, the analysis reduces back to the "easier" case where an irregular arm does not exist.

A bandit adaptation with sublinear regret. Finally, we turn our focus to the learning setting, where the mean payoff functions are initially unknown, and the player has semi-bandit feedback on the realized payoffs of the played arms of each round. In this setting, we begin by analyzing the robustness of our planning algorithm under small perturbations of the payoff functions. Then, we provide sample complexity results to bound the number of samples required to get accurate estimates of the mean payoff functions. By combining the above elements, we transform our planning algorithm into a bandit one, based on an Explore-then-Commit (ETC) scheme, and prove that the latter achieves sublinear regret relative to its planning counterpart. We present these results in Section 5. Due to space constraints, all the omitted proofs of our results have been moved to the Appendix.

2 Preliminaries

Problem definition and notation. We consider a set $\mathcal{A} = \{1, 2, \ldots, n\}$ of n *actions* (or *arms*), where each arm $i \in \mathcal{A}$ is associated with a (mean) *payoff function* $p_i : \mathbb{N} \to [0, 1]$. For each $\tau \in \mathbb{N}$, $p_i(\tau)$ is

the payoff for playing action $i \in \mathcal{A}$, when i has been played before exactly τ rounds (i.e., $\tau = 1$ if the action has been played in the previous round). We refer to the number of rounds since the last play of an arm as the *delay*. For each arm $i \in \mathcal{A}$, we assume that the function $p_i(\tau)$: (a) is monotone non-decreasing in τ, and (b) has a polynomial *recovery time* $\tau_i^{\max}$ such that $p_i(\tau) = p_i(\tau_i^{\max})$ for every $\tau \geq \tau_i^{\max}$. We assume knowledge of a universal upper bound $\tau^{\max}$ on the recovery time of all arms, i.e., $\tau^{\max} \geq \max_{i \in \mathcal{A}} \tau_i^{\max}$.[2] At each round, the player plays a subset of at most $k < n$ actions and collects the sum of their associated payoffs (each computed with respect to the delay of the corresponding action at the time it is played). The objective is to maximize the cumulative payoff collected within an (unknown) horizon of T rounds. We refer to the above instance as k-RB.

As a convention, we assume that the delay of each arm i is initially equal to 1 (i.e., all arms are played at time $t = 0$). Note that, even if we assume that all arms start with delay $\tau^{\max}$ (i.e., the arms have never been played before), our results do not change qualitatively.

For any non-negative integer $n \in \mathbb{N}$, we define $[n] = \{1, 2, \ldots, n\}$. For any vector $\mathbf{w} \in \mathbb{R}^n$ and set $S \subseteq [n]$, we denote $\mathbf{w}(S) = \sum_{i \in S} w_i$. Finally, throughout this work we assume that, when comparing between different payoffs, ties are broken arbitrarily.

Continuous extensions and the correlation gap. Let $f : 2^{[n]} \to [0, \infty)$ be a set function defined over a finite set of n elements. For any vector $\mathbf{y} \in [0, 1]^n$, we denote by $\mathcal{D}(\mathbf{y})$ a distribution over $2^{[n]}$ with marginal probabilities $\mathbf{y} = (y_i)_{i \in [n]}$. We denote by $S \sim \mathcal{D}(\mathbf{y})$ a random subset $S \subseteq [n]$ sampled from the distribution $\mathcal{D}(\mathbf{y})$. Further, we denote by $S \sim \mathcal{I}(\mathbf{y})$ a random subset $S \subseteq [n]$ which is constructed by adding each element $i \in [n]$ to S, independently, with probability y_i.

We recall two canonical continuous extensions of a set function (see [CCPV07, Sch03]):

Definition 2.1 (Multi-linear extension)**.** *For any vector $\mathbf{y} \in [0, 1]^n$, the* multi-linear extension *of a set function f is defined as*

$$F(\mathbf{y}) = \mathop{\mathbb{E}}_{S \sim \mathcal{I}(\mathbf{y})} [f(S)] = \sum_{S \subseteq [n]} f(S) \prod_{i \in S} y_i \prod_{i \notin S} (1 - y_i).$$

Definition 2.2 (Concave closure)**.** *For any vector $\mathbf{y} \in [0, 1]^n$, the* concave closure *of a set function f is defined as*

$$f^+(\mathbf{y}) = \max_{\mathcal{D}(\mathbf{y})} \mathop{\mathbb{E}}_{S \sim \mathcal{D}(\mathbf{y})} [f(S)] = \max_{\alpha} \left\{ \sum_{S \subseteq [n]} \alpha_S f(S) \mid \sum_{S \subseteq [n]} \alpha_S \mathbf{1}_S = \mathbf{y}, \sum_{S \subseteq [n]} \alpha_S = 1, \alpha \succeq 0 \right\},$$

where $\mathbf{1}_S \in \{0, 1\}^n$ is an indicator vector such that $(\mathbf{1}_S)_i = 1$, if $i \in S$, and $(\mathbf{1}_S)_i = 0$, otherwise.

For any non-negative weight vector $\mathbf{w} \in [0, \infty)^n$, of particular importance in the analysis of our algorithm is the function $f_{\mathbf{w}, k} : 2^{[n]} \to [0, \infty)$, defined as

$$f_{\mathbf{w}, k}(S) = \max\{\mathbf{w}(I) \mid I \subseteq S \text{ and } |I| \leq k\}. \tag{1}$$

We remark that $f_{\mathbf{w}, k}$ corresponds to the *weighted rank function* of the (rank-k) uniform matroid under a weight vector $\mathbf{w}$ and, hence, is non-decreasing submodular [Sch03].[3] We denote by $F_{\mathbf{w}, k}$ and $f_{\mathbf{w}, k}^+$ the multi-linear extension and the concave closure of $f_{\mathbf{w}, k}$, respectively.

The *correlation gap* [ADSY10] of a set function is defined as $\sup_{\mathbf{y} \in [0, 1]^n} \frac{f^+(\mathbf{y})}{F(\mathbf{y})}$. The following result due to [Yan11] provides an upper bound on the correlation gap of $f_{\mathbf{w}, k}(\cdot)$:

Lemma 2.3 (Correlation gap [Yan11])**.** *Let $f_{\mathbf{w}, k} : 2^{[n]} \to [0, \infty)$ be the* weighted rank function *of the rank-k uniform matroid. Then, for any vector $\mathbf{y} \in [0, 1]^n$ we have*

$$f_{\mathbf{w}, k}^+(\mathbf{y}) \geq F_{\mathbf{w}, k}(\mathbf{y}) \geq \left(1 - \frac{k^k}{e^k k!}\right) \cdot f_{\mathbf{w}, k}^+(\mathbf{y}) \approx \left(1 - \frac{1}{\sqrt{2\pi k}}\right) \cdot f_{\mathbf{w}, k}^+(\mathbf{y}).$$

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[2]Due to the polynomial recovery time assumption, oracle access to the payoff functions is not required.

[3]A function $f : 2^{[n]} \to \mathbb{R}$ is submodular, if for any $S, T \subseteq [n]$, it holds $f(S \cup T) + f(S \cap T) \leq f(S) + f(T)$.

3 Structural properties of a natural LP relaxation

We consider the following LP relaxation which applies to any planning instance of k-RB:

$$\underset{\mathbf{x} \succeq \mathbf{0}}{\textbf{maximize:}} \quad \sum_{i \in \mathcal{A}} \sum_{\tau \in \mathbb{N}} p_i(\tau) \cdot x_{i,\tau} \qquad (\textbf{LP}_k)$$

$$\textbf{s.t.:} \quad \sum_{i \in \mathcal{A}} \sum_{\tau \in \mathbb{N}} x_{i,\tau} \leq k, \qquad (\text{C.1})$$

$$\sum_{\tau \in \mathbb{N}} \tau \cdot x_{i,\tau} \leq 1, \forall i \in \mathcal{A}. \qquad (\text{C.2})$$

In the above formulation, each variable $x_{i,\tau}$ denotes the fraction of time where arm i is played under delay τ – namely, exactly τ time steps after it was played for the last time. In ($\textbf{LP}_k$), constraint (C.1) follows by the fact that the total fraction of time where any action can be played (under any delay) cannot be more than k. Further, constraints (C.2) are valid constraints for the fractional plays of any individual arm under different delays.

Lemma 3.1. *For any instance of the k-RB problem, let V^* be the optimal value of ($\textbf{LP}_k$) and $\mathrm{OPT}(T)$ be the maximum payoff that can be collected in a time horizon of T rounds. Then $T \cdot V^* \geq \mathrm{OPT}(T)$.*

Assuming knowledge of an upper bound $\tau^{\max} \geq \max_{i \in \mathcal{A}} \{\tau_i^{\max}\}$ on the maximum recovery time of any action, an optimal extreme point solution to ($\textbf{LP}_k$) can be computed in polynomial time:

Fact 3.2. *There exists an optimal solution to ($\textbf{LP}_k$) which is supported in $(x_{i,\tau})_{i \in \mathcal{A}, \tau \in [\tau_i^{\max}]}$. Further, given an upper bound $\tau^{\max}$ on the maximum recovery time of any action, an optimal extreme point solution to ($\textbf{LP}_k$) can be computed in polynomial time.*

Let us introduce the notion of *supported* actions:

Definition 3.3 (Supported actions). *For any (potentially infeasible) solution $\mathbf{x}$ to ($\textbf{LP}_k$), we say that an action $i \in \mathcal{A}$ is supported in $\mathbf{x}$, denoted by $i \in \mathcal{A}(\mathbf{x})$, if there exists $\tau \in \mathbb{N}$ such that $x_{i,\tau} > 0$.*

We introduce the following class of solutions to ($\textbf{LP}_k$):

Definition 3.4 (Delay-feasible solutions). *We say that a vector $\mathbf{x} = (x_{i,\tau})_{i \in \mathcal{A}, \tau \in \mathbb{N}}$ is delay-feasible, if for each supported action $i \in \mathcal{A}(\mathbf{x})$ there exists at most one non-zero variable x_{i,τ_i}, and for this variable it holds $x_{i,\tau_i} = 1/\tau_i$.*

Further, we say that $\mathbf{x}$ is almost-delay-feasible, if it is delay-feasible, possibly with the exception of a single arm $\iota \in \mathcal{A}(\mathbf{x})$, for which there exist at most two non-zero variables $x_{\iota,\tau_{\iota,a}}, x_{\iota,\tau_{\iota,b}}$, such that $x_{\iota,\tau_{\iota,a}} < 1/\tau_{\iota,a}$ and $x_{\iota,\tau_{\iota,b}} < 1/\tau_{\iota,b}$. In the case of almost-delay-feasible solutions, we refer to ι as the irregular arm.

We remark that almost-delay-feasible solutions subsume by definition delay-feasible solutions and that a (almost-)delay-feasible solution is not necessarily a feasible solution to ($\textbf{LP}_k$).

As we prove in the next key-lemma, any extreme point solution of ($\textbf{LP}_k$) is almost-delay-feasible:

Lemma 3.5 (Sparsity pattern of extreme point solutions). *Let $\mathbf{x}$ be any extreme point solution of ($\textbf{LP}_k$). Then, $\mathbf{x}$ is almost-delay-feasible.*

4 Improved approximation guarantees for planning

For any instance of k-RB, our algorithm (see Algorithm 1 below), called "Randomize-Then-Interleave", constructs a feasible planning schedule without requiring knowledge of the time horizon. We remark that for the planning setting of the problem, we treat the payoff of any arm under any possible delay as deterministic.

The algorithm starts from computing an optimal extreme point solution $\mathbf{x}^*$ to ($\textbf{LP}_k$), and then uses this solution to determine a *critical* delay τ_i^* for each supported arm $i \in \mathcal{A}(\mathbf{x}^*)$. By Lemma 3.5, we know that $\mathbf{x}^*$ is almost-delay-feasible, which means that – possibly with the exception of a single irregular arm ι – for every arm $i \in \mathcal{A}(\mathbf{x}^*) \setminus \{\iota\}$, there exists a unique τ_i, such that $x_{i,\tau_i}^* = 1/\tau_i > 0$. Thus, Algorithm 1 sets this unique τ_i to be the critical delay of each arm $i \in \mathcal{A}(\mathbf{x}^*) \setminus \{\iota\}$. In the

case where an irregular arm $\iota \in \mathcal{A}(\mathbf{x}^*)$ exists, then, by Definition 3.4, there exist at most two distinct $\tau_{\iota,a}, \tau_{\iota,b}$ with $x^*_{\iota,\tau_{\iota,a}} > 0$ and $x^*_{\iota,\tau_{\iota,b}} \geq 0$, in which case the critical delay τ^*_ι of ι is set to either $\tau_{\iota,a}$ or $\tau_{\iota,b}$, with marginal probability $\tau_{\iota,a} x^*_{\iota,\tau_{\iota,a}}$ and $\tau_{\iota,b} x^*_{\iota,\tau_{\iota,b}}$, respectively. Notice that by constraints (C.2) of (**LP**$_k$), the above sampling process is well-defined, since $\tau_{\iota,a} x^*_{\iota,\tau_{\iota,a}} + \tau_{\iota,b} x^*_{\iota,\tau_{\iota,b}} \leq 1$. In the case where no τ^*_ι is sampled, the irregular arm is removed from $\mathcal{A}(\mathbf{x}^*)$. After computing the critical delays, for each arm $i \in \mathcal{A}(\mathbf{x}^*)$, the algorithm draws a random *offset* r_i uniformly from $\{0, 1, \ldots, \tau^*_i - 1\}$.

Algorithm 1: Randomize-Then-Interleave

/* Initialization phase   */
Compute an optimal extreme point solution $\mathbf{x}^*$ to (**LP**$_k$).
Let $\iota \in \mathcal{A}$ be the irregular arm (if such an arm exists).
for *each arm $i \in \mathcal{A}(\mathbf{x}^*) \setminus \{\iota\}$* **do**
  | Let τ^*_i be the unique $\tau \in \mathbb{N}$ which satisfies $x^*_{i,\tau} = 1/\tau$.
if *an irregular arm $\iota \in \mathcal{A}(\mathbf{x}^*)$ exists* **then**
  | Let $\tau_{\iota,a}, \tau_{\iota,b} \in \mathbb{N}$ be such that $x^*_{\iota,\tau_{\iota,a}} > 0$, $x^*_{\iota,\tau_{\iota,b}} \geq 0$, and $\tau_{\iota,a} \neq \tau_{\iota,b}$.
  | Set $p_a \leftarrow \tau_{\iota,a} x^*_{\iota,\tau_{\iota,a}}$ and $p_b \leftarrow \tau_{\iota,b} x^*_{\iota,\tau_{\iota,b}}$.
  | Sample τ^*_ι from $\{\tau_{\iota,a}, \tau_{\iota,b}, \infty\}$ with marginals p_a, p_b, and $1 - p_a - p_b$, respectively.
  | **if** $\tau^*_\iota = \infty$ **then**
  | | Remove ι from $\mathcal{A}(\mathbf{x}^*)$.
for *each arm $i \in \mathcal{A}(\mathbf{x}^*)$* **do**
  | Sample an *offset* r_i uniformly at random from $\{0, 1, \ldots, \tau^*_i - 1\}$.
/* Online phase   */
for $t = 1, 2, \ldots$ **do**
  | Let $C_t \subseteq \mathcal{A}$ be the subset of *candidate* arms, defined as $C_t = \{i \in \mathcal{A}(\mathbf{x}^*) \mid t \bmod \tau^*_i \equiv r_i\}$.
  | Play the maximum-payoff subset of k arms in C_t, namely, $A_t = \underset{S \subseteq C_t, |S| \leq k}{\operatorname{argmax}} \sum_{i \in S} p_i(\sigma_{i,t})$,
  |
  | where $\sigma_{i,t}$ (*actual delay*) is the time passed since arm i was last played before t.

After initialization, Algorithm 1 proceeds to its online phase where, at each round t, it first computes a subset of *candidate* arms C_t, namely, all arms $i \in \mathcal{A}(\mathbf{x}^*)$ which satisfy $t \bmod \tau^*_i \equiv r_i$. Then, the algorithm computes the maximum-payoff subset A_t of at most k candidate arms, computed using their actual delays (the number of rounds passed since last play), and plays these arms.

4.1 Approximation analysis for general k

We now present an analysis of the approximation guarantee of Algorithm 1 for general k. In order to facilitate the presentation, we focus on a special class of "easy" instances, where the optimal extreme point solution $\mathbf{x}^*$ of (**LP**$_k$) is delay-feasible (i.e., there is no irregular arm). In Appendix C, we show how this assumption can be relaxed with relatively small additional technical effort.

Recall that we denote by $\mathcal{A}(\mathbf{x}^*)$ the set of arms that are supported in the solution $\mathbf{x}^*$. Notice that, in the case where $\mathbf{x}^*$ is delay-feasible, the only source of randomness in Algorithm 1 is due to the random offsets. Since each offset r_i is sampled independently and uniformly from $\{0, 1, \ldots, \tau^*_i - 1\}$ and $i \in C_t$ if and only if $t \bmod \tau^*_i \equiv r_i$, the following fact is immediate:

Fact 4.1. *Let $\{\tau^*_i\}_{i \in \mathcal{A}(\mathbf{x}^*)}$ be the critical delays of the supported arms. At any round t, each arm $i \in \mathcal{A}(\mathbf{x}^*)$ belongs to C_t independently with probability $1/\tau^*_i$.*

Let us focus on the expected payoff collected by Algorithm 1 at any round $t \geq \tau^{\max}$. By our simplifying assumption that $\mathbf{x}^*$ is delay-feasible, for each supported arm $i \in \mathcal{A}(\mathbf{x}^*)$, there is a unique critical delay τ^*_i, such that $x^*_{i,\tau^*_i} = 1/\tau^*_i$ is the only non-zero variable of $\mathbf{x}^*$ corresponding to i.

We construct a vector $\mathbf{y} = [0,1]^n$, such that $y_i = 1/\tau^*_i$ for each $i \in \mathcal{A}(\mathbf{x}^*)$, and $y_i = 0$, otherwise. Further, let us define a vector $\mathbf{w} \in [0, \infty)^n$ such that $w_i = p_i(\tau^*_i)$ for each $i \in \mathcal{A}(\mathbf{x}^*)$, and $w_i = 0$, otherwise. Recall that for any fixed round $t \geq \tau^{\max}$, by definition of C_t and the construction of the offsets, the events $\{i \in C_t\}_{i \in \mathcal{A}}$ are independent. Hence, we can think of C_t as a random set, where each arm $i \in \mathcal{A}$ belongs to C_t independently with probability y_i, namely, $C_t \sim \mathcal{I}(\mathbf{y})$.

Let $\underline{A}_t = \underset{S \subseteq C_t, |S| \le k}{\arg\max} \sum_{i \in S} p_i(\tau_i^*)$ be a set of at most k arms in C_t with maximum total payoff measured with respect to the critical delays. By construction of C_t, the actual delay $\sigma_{i,t}$ of each $i \in \mathcal{A}(\mathbf{x}^*)$ is (deterministically) lower bounded by τ_i^*, since any arm must be a candidate in order to be played (which happens exactly every τ_i^* rounds). Thus, for the expected collected payoff we have

$$\underset{C_t \sim \mathcal{I}(\mathbf{y})}{\mathbb{E}} \left[\max_{S \subseteq C_t, |S| \le k} \sum_{i \in S} p_i(\sigma_{i,t}) \right] \ge \underset{C_t \sim \mathcal{I}(\mathbf{y})}{\mathbb{E}} \left[\max_{S \subseteq C_t, |S| \le k} \sum_{i \in S} p_i(\tau_i^*) \right] = \underset{C_t \sim \mathcal{I}(\mathbf{y})}{\mathbb{E}} \left[\mathbf{w}(\underline{A}_t) \right]. \quad (2)$$

Using Equation (1) in combination with Definitions 2.1 and 2.2, we get

$$\underset{C_t \sim \mathcal{I}(\mathbf{y})}{\mathbb{E}} \left[\mathbf{w}(\underline{A}_t) \right] = \underset{C_t \sim \mathcal{I}(\mathbf{y})}{\mathbb{E}} \left[f_{\mathbf{w},k}(C_t) \right] = F_{\mathbf{w},k}(\mathbf{y}) \ge \left(1 - \frac{k^k}{e^k k!} \right) \cdot f_{\mathbf{w},k}^+(\mathbf{y}), \quad (3)$$

where the last inequality follows by Lemma 2.3.

The following technical lemma relates $f_{\mathbf{w},k}^+(\mathbf{y})$ with optimal value of ($\mathbf{LP}_k$):

Lemma 4.2. *For any vectors $\mathbf{w} \in [0, \infty)^n$ and $\mathbf{y} \in [0, 1]^n$, let $f_{\mathbf{w},k}^+(\mathbf{y})$ be the concave closure of the weighted rank function of the rank-k uniform matroid with weights $\mathbf{w}$, evaluated at $\mathbf{y}$. Then, if $\mathbf{y}$ satisfies $\| \mathbf{y} \|_1 \le k$, it is the case that $f_{\mathbf{w},k}^+(\mathbf{y}) \ge \sum_{i \in [n]} w_i y_i$.*

Let V^* be the optimal value of ($\mathbf{LP}_k$). Since $\| \mathbf{y} \|_1 \le k$, then by Lemma 4.2, we have that

$$f_{\mathbf{w},k}^+(\mathbf{y}) \ge \sum_{i \in [n]} w_i \cdot y_i = \sum_{i \in \mathcal{A}(\mathbf{x}^*)} p_i(\tau_i^*) \cdot x_{i,\tau_i^*}^* = V^*. \quad (4)$$

By combining Equations (2) to (4) with Lemma 3.1, it follows that, for any round $t \ge \tau^{\max}$, the expected payoff collected by Algorithm 1 is at least $\left(1 - k^k/e^k k!\right)$-times the average optimal payoff. By using linearity of expectations together with the fact that all payoffs are bounded in $[0, 1]$, this suffices to prove a long-run $\left(1 - k^k/e^k k!\right)$-approximation guarantee (in expectation) for instances where $\mathbf{x}^*$ is delay-feasible.

In Appendix C, we show how the above analysis can be extended for the general case where $\mathbf{x}^*$ is almost-delay-feasible (i.e., an irregular arm exists). The main idea is to repeat the above steps, *conditioned* on any possible realized critical delay of the irregular arm, and then combine the resulting guarantees using total probability. However, conditioning on the realized critical delay of the irregular arm causes a subtle technical issue: the resulting (delay-feasible) solution *after the rounding* can be infeasible for ($\mathbf{LP}_k$) and, hence, the condition of Lemma 4.2 that $\| \mathbf{y} \|_1 \le k$ may no longer be true. As we show in the proof of the following theorem, tackling this issue comes at a cost of only an extra $k/k+1$-factor in the approximation guarantee.

Theorem 4.3. *For any instance of k-RB, where $\mathrm{OPT}(T)$ is the optimal expected payoff that can be collected in T rounds, the total expected total payoff collected by Algorithm 1 is at least*

$$\frac{k}{k+1} \cdot \left(1 - \frac{k^k}{e^k k!} \right) \mathrm{OPT}(T) - \mathcal{O}(k \cdot \tau^{\max}) \quad \approx \quad \left(1 - \mathcal{O}\left(\frac{1}{\sqrt{k}} \right) \right) \mathrm{OPT}(T) - \mathcal{O}(k \cdot \tau^{\max}).$$

We remark.that, although the asymptotic approximation guarantee of Algorithm 1 is the same as the one in [SLZZ21], the constants of our algorithm are significantly better. In particular, for $k \in \{2, 3, 4, 5, 10\}$, the guarantees of [SLZZ21] and Algorithm 1 are $\{0.33, 0.40, 0.44, 0.46, 0.57\}$ and $\{0.48, 0.58, 0.64, 0.68, 0.79\}$, respectively.

4.2 Alternative and tighter analysis for $k = 1$

We now provide an alternative analysis for the case of 1-RB, in which we manage to avoid the extra $k/k+1$-multiplicative loss in the approximation guarantee – thus bringing the state-of-the-art in this case to $1 - 1/e$. As we have already discussed in Section 4.1 and the proof of Theorem 4.3, this extra loss is due to the existence of an irregular arm.

Let $\mathbf{x}^*$ be an optimal solution to ($\mathbf{LP}_k$) for $k = 1$ and let $\iota \in \mathcal{A}(\mathbf{x}^*)$ be the irregular arm with $\tau_{\iota,a}^*$ and $\tau_{\iota,b}^*$ the two potential critical delays such that $x_{\iota,\tau_{\iota,a}}^*, x_{\iota,\tau_{\iota,b}}^* \ge 0$ (with at least one of $x_{\iota,\tau_{\iota,a}}^*$ and

$x^*_{\iota, \tau_{\iota, b}}$ being strictly positive, by Definition 3.4). Instead of conditioning on the sampled critical delay of the irregular arm ι, we construct two parallel copies ι_a and ι_b – one for each possible realization of its critical delay (and associated payoff). Let us define $\mathcal{A}' = \mathcal{A}(\mathbf{x}^*) \setminus \{\iota\} \cup \{\iota_a, \iota_b\}$ to be the supported arms of $\mathbf{x}^*$, where we replace the irregular arm ι with the two parallel copies ι_a and ι_b.

In the next lemma, we characterize the marginal probability that any arm in $\mathcal{A}'$ is added to the set of candidate arms at any round. Notice that, as opposed to Fact 4.1, the randomness in this case is taken over both the sampling of a critical delay for the irregular arm as well as the random offsets.

Lemma 4.4. *For any arm $i \in \mathcal{A}'$ and at any round t, we have that $\mathbb{P}\left[i \in C_t \text{ and } \tau_i^* = \tau\right] = x^*_{i,\tau}$.*

Notice that, for the parallel copies of the irregular arm, the events $\{\iota_a \in C_t\}$ and $\{\iota_b \in C_t\}$ are now dependent and, in particular, *mutually exclusive* (since, eventually, only one of them is realized). However, for the special case where $k = 1$, we show that the expected payoff collected at any round $t \geq \tau^{\max}$ can only decrease if the two events were instead independent with the same marginals.

Let us relabel the arms in $\mathcal{A}'$ in non-increasing order of $p_i(\tau_i^*)$ (breaking ties arbitrarily). Note that by exchanging ι with the two copies ι_a and ι_b the cardinality of $\mathcal{A}'$ can be at most $n + 1$.

Let us construct a $(n+1)$-dimensional vector $\mathbf{y} \in [0, 1]^{n+1}$, such that $y_i = x^*_{i,\tau_i^*}$ for any $i \in \mathcal{A}'$, and $y_i = 0$, otherwise. Further, we construct a weight vector $\mathbf{w} \in [0, \infty)^{n+1}$, such that $w_i = p_i(\tau_i^*)$ for any $i \in \mathcal{A}'$, and $w_i = 0$, otherwise. Let $\mathcal{D}(\mathbf{y})$ be a distribution over subsets of $\mathcal{A}'$, where every arm $i \in \mathcal{A}' \setminus \{\iota_a, \iota_b\}$ (except for ι_a and ι_b) is added to a set $S \sim \mathcal{D}(\mathbf{y})$, independently, with probability x^*_{i,τ_i^*}. The arms ι_a and $\iota_b \in \mathcal{A}'$ are also added to $S \sim \mathcal{D}(\mathbf{y})$ with marginal probabilities $x^*_{\iota_a, \tau^*_{\iota_a}}$ and $x^*_{\iota_b, \tau^*_{\iota_b}}$, respectively, yet not independently (since they are mutually exclusive). Therefore, by the above definitions and Lemma 4.4, the set of candidate arms at any fixed round $t \geq \tau^{\max}$ is distributed as $C_t \sim \mathcal{D}(\mathbf{y})$.

Let us fix any round $t \geq \tau^{\max}$. By using the fact that the actual delay $\sigma_{i,\tau}$ of each arm $i \in \mathcal{A}'$ is deterministically lower bounded by its critical delay τ_i^*, we can lower bound the expected payoff collected by our algorithm at time t (for the case where $k = 1$) as

$$\mathop{\mathbb{E}}_{C_t \sim \mathcal{D}(\mathbf{y})}\left[\max_{i \in C_t} p_i(\sigma_{i,t})\right] \geq \mathop{\mathbb{E}}_{C_t \sim \mathcal{D}(\mathbf{y})}\left[\max_{i \in C_t} p_i(\tau_i^*)\right].$$

As we show in the following lemma, the RHS of the above inequality can only decrease, if we replace $C_t \sim \mathcal{D}(\mathbf{y})$ with $C_t \sim \mathcal{I}(\mathbf{y})$, namely, if we assume that the two copies ι_a and ι_b are added to C_t independently, but with the same marginals.

Lemma 4.5. *For any round t and distributions $\mathcal{D}(\mathbf{y})$ and $\mathcal{I}(\mathbf{y})$ defined as above, we have that*

$$\mathop{\mathbb{E}}_{C_t \sim \mathcal{D}(\mathbf{y})}\left[\max_{i \in C_t} p_i(\tau_i^*)\right] \geq \mathop{\mathbb{E}}_{C_t \sim \mathcal{I}(\mathbf{y})}\left[\max_{i \in C_t} p_i(\tau_i^*)\right].$$

The intuition behind the above result is the following: fix any two mutually exclusive events X and Y drawn from a joint distribution $(X, Y) \sim \mathcal{D}'$, with marginal probabilities $\mathbb{P}[X] > 0$ and $\mathbb{P}[Y] > 0$, respectively. It can be trivially verified that the probability of their union, $\mathbb{P}[X \cup Y] = \mathbb{P}[X] + \mathbb{P}[Y]$, is strictly larger than the corresponding probability of any other joint distribution (including the product distribution), where the two events are not mutually exclusive yet maintain the same marginals. However, for the case where $k = 1$, since at most one arm can be played at each round, the contribution of ι_a and ι_b to the expected payoff is only affected by $\mathop{\mathbb{P}}_{C_t \sim \mathcal{D}(\mathbf{y})}\left[\{\iota_a \in C_t\} \cup \{\iota_b \in C_t\}\right]$.

After establishing the above fact, the rest of the proof then follows exactly as the proof of Theorem 4.3.

Theorem 4.6. *For any instance of 1-RB, the expected total payoff collected by Algorithm 1 is at least*

$$(1 - 1/e)\,\mathrm{OPT}(T) - \mathcal{O}(\tau^{\max}).$$

5   Online learning with sublinear regret

In this section, we consider the learning setting of k-RB where, for every arm $i \in \mathcal{A}$ and delay τ, the payoff for playing arm i under delay τ is drawn independently from a distribution of mean $p_i(\tau)$, bounded in $[0, 1]$. Assuming semi-bandit feedback on the realized payoffs, our goal is to show that

there exists a bandit variant of Algorithm 1 with sublinear γ_k-approximate regret guarantees, where γ_k is the (multiplicative) approximation guarantee of Algorithm 1 (provided in Theorems 4.3 and 4.6) for an instance of k-RB.

Robustness of Algorithm 1. As a first step, we study the robustness of Algorithm 1 under small perturbations of the payoff functions; in particular, we analyze the approximation guarantee of the algorithm when it runs using estimates $\{\widehat{p}_i\}_{i\in\mathcal{A}}$ instead of the actual payoffs $\{p_i\}_{i\in\mathcal{A}}$. We remark that the monotonicity of the (estimated) payoff functions is *not required* for the correctness of Algorithm 1, but only for proving its approximation guarantee.

Lemma 5.1. *Let $\widehat{p}_i$ be an estimate of the payoff function p_i for each $i \in \mathcal{A}$, such that $|\widehat{p}_i(\tau) - p_i(\tau)| \leq \epsilon, \forall \tau \in \mathbb{N}$ for some $\epsilon \in (0,1)$. The expected payoff collected in T rounds by Algorithm 1, when it uses the estimates instead of the actual payoff functions, is lower bounded by*

$$\gamma_k \operatorname{OPT}(T) - \mathcal{O}(k \cdot \epsilon \cdot T + k \cdot \tau^{\max}).$$

Estimation of the payoff functions. By using standard concentration results [Hoe63], we bound the number of samples required for each arm-delay pair in order to get an ϵ-estimate $\widehat{p}_i$ for the mean payoff function p_i of every arm $i \in \mathcal{A}$.

Lemma 5.2. *For every $\epsilon, \delta \in (0,1)$, let $\widehat{p}_i(\tau)$ be the empirical average of samples drawn from $p_i(\tau)$ for every arm $i \in \mathcal{A}$ and delay τ, using $m = \frac{1}{2\epsilon^2} \ln\left(\frac{2\tau^{\max}n}{\delta}\right)$ samples, where n is the number of arms. Then, with probability $1 - \delta$, it holds $|\widehat{p}_i(\tau) - p_i(\tau)| \leq \epsilon$ for every $i \in \mathcal{A}$ and $\tau \in [\tau^{\max}]$.*

The following observation provides an upper bound on the number of rounds required in order to get m samples for each arm-delay pair in any k-RB instance:

Fact 5.3. *In any k-RB instance, the number of rounds required to collect m independent samples from $p_i(\tau)$, for each arm $i \in \mathcal{A}$ and delay $\tau \in [\tau^{\max}]$, is upper bounded by $\mathcal{O}(nm(\tau^{\max})^2/k)$.*

Bandit algorithm with sublinear regret. By combining the above elements, it is easy to design an Explore-then-Commit (ETC) variant of Algorithm 1, proving that the algorithm accepts sublinear regret in the bandit setting. For simplicity, we assume here that the time horizon T is known to the player – an assumption that can be relaxed using the doubling trick method [LS20].

Let $m = 1/2\epsilon^2 \ln(2\tau^{\max}n/\delta)$ for some $\epsilon, \delta \in (0,1)$. In the first $\mathcal{O}(nm(\tau^{\max})^2/k)$ rounds, the bandit algorithm collects m independent samples from $p_i(\tau)$, for each $i \in \mathcal{A}$ and $\tau \in [\tau^{\max}]$. Then, by taking the empirical average of these samples, it constructs estimates $\widehat{p}_i(\tau)$ for each i and τ. For the rest of the time horizon, the algorithm simulates Algorithm 1, using the empirical estimates $\widehat{p}_i(\tau)$ in place of the unknown payoff functions.

By setting the parameters ϵ, δ accordingly and combining Lemmas 5.1 and 5.2 with Fact 5.3, we prove the following result:

Theorem 5.4. *There exists a learning adaptation of Algorithm 1 in the semi-bandit setting, for which the γ_k-approximate regret for T rounds can be upper bounded as*

$$\mathcal{O}\left((k \cdot n)^{\frac{1}{3}}(\tau^{\max})^{\frac{2}{3}}T^{\frac{2}{3}}\ln(\tau^{\max}nT) + k \cdot \tau^{\max}\right).$$

Conclusion and further directions

Following the recent line of work on non-stationary bandits with recharging payoffs, we revisited one of the most general formulations studied in the area, and significantly improved the state-of-the-art for the planning setting. In particular, when at most one arm is played per round, we designed an algorithm that collects at least 63% (asymptotically and in expectation) of the optimal payoff, improving the best-known 25%, due to prior work. By providing results on the robustness and sample complexity of our algorithm, we transformed the latter into a bandit algorithm with sublinear regret. Our work leaves a number of interesting open questions. An approximation guarantee of $(1 - 1/e)$ has been proved to be the best achievable – under standard complexity assumptions – for a number combinatorial problems. Therefore, an immediate future direction would be to either provide a hardness result for the planning setting of the problem, or to further improve the best-known approximation. Another interesting question would be to explore whether a one can design a bandit adaptation of our algorithm with improved regret guarantees. Finally, the empirical evaluation of our algorithm on real (or artificial) data is another interesting direction.

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