A Likelihood Ratio Test of Genetic Relationship among Languages

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Abstract

Lexical resemblances among a group of languages indicate that the languages could be genetically related, i.e., they could have descended from a common ancestral language. However, such resemblances can arise by chance and, hence, need not always imply an underlying genetic relationship. Many tests of significance based on permutation of wordlists and word similarity measures appeared in the past to determine the statistical significance of such relationships. We demonstrate that although existing tests may work well for bilateral comparisons, i.e., on pairs of languages, they are either infeasible by design or are prone to yield false positives when applied to groups of languages or language families. To this end, inspired by molecular phylogenetics, we propose a likelihood ratio test to determine if given languages are related based on the proportion of invariant character sites in the aligned wordlists applied during tree inference. Further, we evaluate some language families and show that the proposed test solves the problem of false positives. Finally, we demonstrate that the test supports the existence of macro language families such as Nostratic and Macro-Mayan.

1 Introduction

Languages that descend from a common ancestral language are termed to be *genetically related*. The existence of lexical resemblances between the two languages is a preliminary indication that they could be related. Such resembling lexicons that truly have a common origin are called *cognates*. For instance, Sanskrit *nāma* and English *name* are cognates that can be traced to Proto-Indo-European **h₃nómn*. However, such resemblances can occur accidentally or out of sheer chance. For instance, Persian *bad* and *behtar* accidentally resemble English *bad* and *better* respectively, and are not true cognates. Hence, it is necessary to show *statistically significance* on any appropriate measure that captures the lexical relatedness before arguing for a genetic relationship among any group of languages or language families (Campbell, 2013).

Several significance tests appeared in the past to address this problem, with the majority of them based on permutation tests, starting from Oswalt (1970). Given wordlists of a group of languages to be evaluated for a genetic relationship, these tests obtain the null distribution of a certain measure capturing similarity between word pairs by random permutations of the wordlists. Such tests either act *bilaterally*, i.e., on a pair of languages or proto-languages, or *multilaterally* on a group of languages. Among these, the multilateral comparison, as made famous by Greenberg (1963, 1971, 1987, 2000) in traditional historical linguistics, has been a subject of much criticism (Poser and Campbell, 2008). Hence, the preferred way of comparing two language families has been to compare their reconstructed proto-forms bilaterally. However, Greenberg (2005) argues that genetic classification should precede proto-language reconstruction. Moreover, there is often a lack of agreement on reconstructed proto-forms both in terms of phonology and semantics which gives room for sufficient manipulation of wordlists that can in turn alter the results of significance tests (Kessler, 2015). Further, we demonstrate that multilateral permutation tests (Kessler and Lehtonen, 2006; Kessler, 2007) yield false negatives even after incorporating complex word similarity metrics such SCA and LexStat (List, 2010, 2012).

To overcome these issues, we turn to *phylogenetic analysis* (Wiley and Lieberman, 2011) that is known to approximately capture the ancestral states and has been applied to phonological reconstruction tasks such as proto-language and cognate reflex prediction tasks (Jäger, 2019, 2022) with reasonably good results. Specifically, we propose a *likelihood ratio test* (LRT) where we expect the difference in likelihoods of the best trees to

capture the compatibility of a group of languages to form a language family assuming two different invariant (character) site proportions for null and alternate hypotheses. Intuitively, related languages should have more positions than unrelated languages where a character or a sound class is invariant. Hence, we essentially capture the notion of relatedness relatively as possessing a higher proportion of invariant sites. Further in this test, reconstructed proto-forms are not required and at the same time, the evolutionary tree structure is strictly imposed by design, unlike the multilateral model, thereby effectively circumventing the aforementioned methodological problems. Although inspired by similar tests from molecular phylogenetics, the test we propose is novel in the sense that the problem of testing common descent never arises in biology since monogenesis is accepted as a fact therein (Kessler, 2008). We further evaluate the test on various language families and demonstrate that the test does not misclassify unrelated languages as related. We finally show that the test supports the existence of the macro-families Nostratic (Bomhard and Kerns, 1994) and Macro-Mayan (Campbell, 1997). While such an attempt to justify the existence of macro-families using bootstrap analysis of distance-based phylogeny is found in Jäger (2015), preference to express statistical significance in terms of likelihood ratio over bootstrap support values whose interpretation is debated in molecular phylogenetics (Anisimova and Gascuel, 2006).Our contributions are summarized as follows.

- We have proposed a likelihood ratio test to determine the genetic relatedness of a group of languages based on invariant site proportions.
- We have demonstrated by applying various language sets that the test does not exhibit the problem of false positives nor requires reconstructed proto-forms, unlike the previously proposed tests.
- We have found through the test some supporting evidence for the existence of macro-families namely Nostratic and Macro-Mayan

The rest of the paper is summarized as follows. Related work is discussed in §2. The methodology of the test is presented in §3. Evaluation details such as datasets and the details of previous methods, and variants are discussed in §4. The results of evaluation along with application on long-range comparisons are discussed in §5. The paper is concluded in §7.

2 Related Work

Notable significance tests for language relationships are discussed as follows. Permutation test for bilateral comparisons was introduced by Oswalt (1970). The significance of sound correspondences by brute force probability calculation was proposed by Ringe (1992, 1996). This approach was however criticized for not being able to show significance for known related pairs of languages like Latin-English and also for accounting phonologically implausible sound correspondences (Kessler, 2001). Multilateral permutation tests were proposed by Kessler and Lehtonen (2006); Kessler (2007). Several applications of permutation tests exist such as Turchin et al. (2010), Kassian et al. (2015) etc.

Some notable likelihood ratio tests in molecular phylogenetics mostly on topologies include Huelsenbeck and Bull (1996), Huelsenbeck et al. (1996), Goldman et al. (2000), Anisimova and Gascuel (2006) etc. where basic bootstrap analysis is argued to be insufficient to establish statistical significance on phylogenies. Otherwise, support for macro-families through bootstrap analysis for distance-based trees is shown in Jäger (2015). Comparisons of various methods of phylogenetic reconstruction such as distance-based, binary-character-based, etc. are given by Jäger (2018). Sound-class character-based phylogenetic analysis is found in Jäger (2019) and Jäger (2022). Best phylogenetic reconstructions are however achieved by Bayesian inference on binary character-based cognate encodings (Rama et al., 2018; Rama and List, 2019).

3 Methodology

The key concept revolves around the idea that any hypothesis, in this case, a hypothesis on a phylogeny, is preferred over a competing null hypothesis if it is significantly more likely, i.e., has a higher likelihood than the latter. Given the wordlist data encoded as an aligned character matrix, related languages are expected to have a higher number of *invariant* columns. Thus, our null hypothesis consists of a phylogeny with a small proportion (fixed at 1%) of invariant sites, whereas the alternative hypothesis consists of a phylogeny with a larger but reasonable proportion (fixed at 6%) of invariant sites. The observed difference in their likelihood

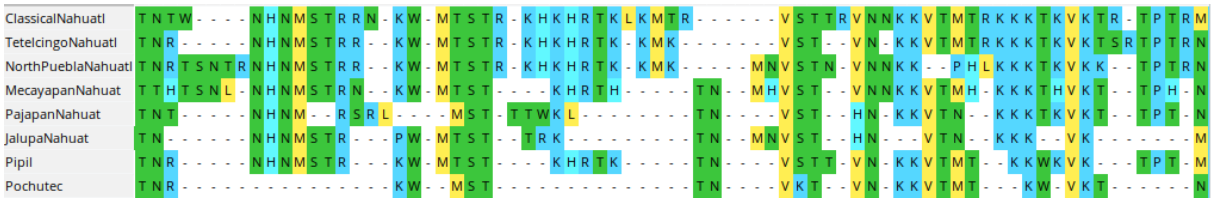


Figure 1: A section of character matrix for Uto-Aztecan family consisting of concatenated MSAs of consonant classes, one from each concept

of real data is compared with that of data simulated from the null hypothesis through parametric bootstrapping and, accordingly, one of the hypotheses is rejected. The steps are elaborated next.

3.1 Character Matrix

The wordlists of a given group of languages, as mentioned previously, are encoded in the form of a character matrix. It consists of concatenated aligned words per concept, i.e., meaning. Thus, each row represents a language or *taxon*, and each column, also referred to as *site* in this paper, consists of phoneme classes, e,g., Dolgopolsky classes. Formally, let the input language set be $\{L_1, \ldots, L_m\}$, whose genetic relatedness is to be verified statistically. Let there be n concepts $C_1, \ldots C_n$ in the wordlists. Each language L_i should have for each concept C_j a single word, say w_{ij}. If a language has multiple words for a single semantic slot, only the one with fundamental meaning is retained. In case the language has no word for a semantic slot, it is represented as a gap '–'. For each concept C_j and alphabet set $\mathbb{A}$, let $W^j \in \mathbb{A}^{m \times l_j}$ represent a multiple sequence alignment (MSA) of words where l_j is the length or the number of phonemes (with vowels removed) in each word. The final character matrix $X \in \mathbb{A}^{m \times N}$ is concatenation of W^j, i.e., $[W^1 \ldots W^n]$ across columns and $N = \sum_{j=1}^{n} l_j$. For an illustration of a final character matrix, see Fig. 1 (generated by MEGA11 (Tamura et al., 2021)).

3.2 Substitution Model

A substitution model describes the evolution of a character at a site assuming a Markovian process. Various substitution models have been described for various alphabets, i.e., for nucleotides, amino acids, etc. In this paper, we assume the simplest possible model where substitution rates are assumed to be equal between all the pairs of distinct characters. The resultant model is known as the Jukes-Cantor model (Jukes et al., 1969) in case of nucleotide substitutions and as Poisson (Bishop and Friday, 1987) in case of amino-acid substitutions. Formally, let the number of characters in the alphabet $\mathbb{A}$ be N. An element q_{ij} of the rate matrix Q, which denotes the rate at which character i mutates to j is defined as follows:

$$q_{ij} = \mu \cdot \pi_i \,,\, i \neq j \text{ (equal rates)} \tag{1}$$

where π_i denotes the frequency of character i at the site. The diagonal element should satisfy the normalization constraint:

$$q_{ii} = -\sum_{j \neq i} q_{ij} \tag{2}$$

The probability of transition $i \rightarrow j$ in time t is given by the matrix $P(t) = \{p_{ij}\} = e^{Qt}$. Likelihood of an evolutionary tree with topology T can be, thus, calculated from the substitution matrix where branch lengths V would denote the time.

3.3 Maximum Likelihood Tree

For any phylogenetic tree with topology T, branch lengths V, other parameters such as shape parameter of heterogeneous rate, proportion of invariant sites, etc. denoted by Θ, and with the observed data i.e., character matrix X, the *likelihood* is defined as the product of likelihoods at each site assuming independence:

$$\mathcal{L}(T, V, \Theta | X) = \prod_{i=1}^{N} P(X_i | T, V, \Theta) \tag{3}$$

The parameters that maximize the likelihood, $\hat{T}, \hat{V}$, and $\hat{\Theta}$, define the *maximum likelihood tree* which is usually obtained by some heuristic search in the parameter space.

3.4 Invariant Sites

Invariant sites are those sites that are constant or evolve very slowly. These can be estimated through a maximum likelihood search along with other parameters. The proportion of invariant sites, P_{inv}

may be known beforehand or estimated. Given the invariant sites, the likelihood defined in §3.3 is only the product of likelihoods across the variant sites.

Our observation is that estimated P_{inv} is higher (>0.06) among related languages while lower (≈0.01) among (possibly) unrelated languages. Based on this observation and preliminaries, we now describe the likelihood ratio test.

3.5 Likelihood Ratio Test (LRT)

Given a null hypothesis H_0 and a competing alternative hypothesis H_a, the latter is preferred if it is more likely than the former i.e., $\mathcal{L}_{H_a} > \mathcal{L}_{H_0}$. In our case, the hypotheses consist of respective phylogenetic tree parameters estimated for best trees i.e., H_0 consists of $\hat{T}_0, \hat{V}_0, \hat{\Theta}_0$ and H_a consists of $\hat{T}_a, \hat{V}_a, \hat{\Theta}_a$. The likelihood ratio test defines the following metric to decide whether to reject the null hypothesis:

$$\delta = 2 \cdot \ln \left(\frac{\mathcal{L}(\hat{T}_a, \hat{V}_a, \hat{\Theta}_a)}{\mathcal{L}(\hat{T}_0, \hat{V}_0, \hat{\Theta}_0)} \right) \quad (4)$$

The Likelihood Ratio Test (LRT) metric δ was shown to asymptotically follow a chi-squared distribution when the null hypothesis is assumed with the degrees of freedom $p - q$, where p and q respectively are the numbers of free parameters in the alternate and the null hypotheses (Wilks, 1938). However, it was argued that this may not hold in general for phylogenetic problems due to the discrete nature of tree topology (see Huelsenbeck and Bull (1996); Huelsenbeck et al. (1996); Anisimova and Gascuel (2006) for relevant work). As a result, the distribution of δ is determined by a parametric bootstrapping method where it is measured on the data simulated by the parameters estimated assuming the null hypothesis H_0, i.e, using the parameters $\hat{T}_0$, $\hat{V}_0$ and $\hat{\Theta}_0$.

As mentioned in §3.4, we propose LRT to test the relatedness of a group of languages using varying proportions of invariant sites. In other words words the null hypothesis H_0 consists of invariant site proportion P^0_{inv} and alternate hypothesis H_a consists of P^a_{inv} where $P^0_{inv} < P^a_{inv}$ as per the observations discussed in §3.4.

The typical way of obtaining the distribution for δ under H_0 involves finding the parameters $\{\hat{T}_0, \hat{V}_0, \hat{\Theta}_0\}$ and $\{\hat{T}_a, \hat{V}_a, \hat{\Theta}_a\}$ for the best trees respectively under H_0 and H_a along with observed δ, say $\hat{\delta}$. Further, generating several, say k, bootstrap replicates from the topology, branch lengths, and other parameters defined by $\{\hat{T}_0, \hat{V}_0, \hat{\Theta}_0\}$, i.e., assuming H_0. Finally, by running the maximum likelihood search again on these replicated to obtain several samples for δ, say $\{\delta_1, \ldots, \delta_k\}$. However, we found considerable variation in $\hat{\delta}$, since the maximum likelihood search is only a heuristic and is affected by initialization. As a result, we obtain several samples for $\hat{\delta}$, say $\{\hat{\delta}_1, \ldots, \hat{\delta}_k\}$ by running the search k times and based on the null parameters, a single bootstrap replicate is generated for each search to consequently obtain $\{\delta_1, \ldots, \delta_k\}$ for corresponding k searches. Finally the *p-value* for $\mathbb{E}[\delta] < \mathbb{E}[\hat{\delta}]$ is obtained by one-sided paired t-test.

Family	Languages	Concepts	Words
Afrasian (AfA)	21	39	770
Lolo-Burmese (LoBur)	15	39	565
Mon-Khmer (MKh)	9	199	1701
Munda (Mun)	4	199	759
Mon-Khmer (MKh)	16	94	1332
Mayan (May)	30	94	2667
Mixe-Zoque (MZ)	10	94	905
Uto-Aztecan (UAz)	9	94	803
Indo-European (IE)	12	185	2209
Dravidian (Drav)	4	183	716
Kartvelian (Kart)	1	180	180

Table 1: Metadata of the language families considered in this study

Family	Languages	Concepts	Words
Austro-Asiatic (AA)	58	200	11001
Austronesian (AN)	45	210	8309
Indo-European (IE)	42	208	8478
Pama-Nyungan (PN)	67	183	11503
Sino-Tibetan (ST)	64	110	6762

Table 2: Metadata of the dataset for tree construction

4 Experimental Setup

The section discusses the details of the experiments including datasets, baseline models, and implementation details as follows.

4.1 Datasets

The data for evaluating the tests consists of wordlists from multiple language (sub-)families and their combinations. Combinations of related sub-families serve as positive examples while those of unrelated serve as negative examples. Evaluating the macro-families also consists of language groups whose relationship is only distantly suggested such as Nostratic (Bomhard and Kerns, 1994).

The details of data from each family are shown in Table 1. Out of these, Mon-Khmer and Munda (200

wordlists) are extracted from the Austro-Asiatic data from Rama et al. (2018). Data for Old languages of Nostratic comprising Indo-European, Dravidian, and Kartvelian are prepared by us from the Swadesh 200-wordlists available at Wiktionary[1]. Data for all the other families are obtained from Rama (2018) which were, in turn, collected from various publicly available sources.

In the Nostratic grouping, we considered the languages that are surviving or have surviving descendants and were attested latest by the 10th centuary CE. The motivation behind this choice is that older languages should be closer to the ancestral language and each other if at all there is any relationship. Several languages including literary Dravidian languages, Georgian, Armenian, etc. are mostly conservative and deviate a little from their old forms. The data is pre-processed by excluding motivated word forms including onomatopoeia, nursery forms, etc. listed in Kessler (2001). Short forms, i.e., words consisting of single syllables are also excluded. Such cleaning is necessary to avoid the appearance of spurious relationships. In the case of Nostratic, we were also careful to exclude borrowals by tracing etymologies from Wiktionary. This step could not be extended to other language families due to lack of readily available etymological information.

All the methods employed in this work, including both the proposed one and baseline ones described in §4.2, involve the construction of a phylogenetic tree. Hence, we also compare the methods on a tree construction task where we see how well the trees match the golden truth trees wherever available. The data for this task is taken from Rama et al. (2018) as summarized in Table 2.

4.2 Multilateral Permutation Test

As mentioned in §1, most previous methods compare languages bilaterally, i.e., a pair at a time. As a result, the only possible way to compare the language families in this approach is to compare their reconstructed proto-languages. However, proto-forms of a proto-language are not often universally agreed which leads to considerable allowance of manipulation that can affect the results (Kessler, 2015). An alternate solution to determine the significance of relationship among multiple languages was proposed by Kessler and Lehtonen (2006); Kessler (2007) who employ a permutation test based on multilateral comparison. This has been well received in historical linguistics (Ringe and Eska, 2013).

The test is based on nearest-neighbour hierarchical clustering where at any point two closest clusters are lumped into one cluster. The basic distance measure, $\hat{d}(A, B)$, between any two clusters A and B is the average of distances between all possible pairs of languages in these clusters viz.,

$$\hat{d}(A, B) = \frac{1}{|A| \cdot |B|} \sum_{a \in A} \sum_{b \in B} d(a, b) \qquad (5)$$

where the distance $d(a, b)$ between any two languages a and b is the mean distance between the pairs of words over all concepts. Following the notations of §3.1,

$$d(a, b) = \frac{\sum_{C_j, w_{aj} \neq \emptyset, w_{bj} \neq \emptyset} d(w_{aj}, w_{bj})}{|\{C_j : w_{aj} \neq \emptyset, w_{bj} \neq \emptyset\}|} \qquad (6)$$

Considering mean essentially enforces multilateral comparison. Alternatively, this is the same as UPGMA (Sokal and Michener, 1958) tree construction where at any bifurcating node, a uniform rate of evolution is assumed across daughter clades. The final similarity metric $\hat{s}(A, B)$ is determined by the following statistic that is computed based on a random permutation of words across each column (taxon) which yields random distances $d(A, B)$:

$$\hat{s}(A, B) = \frac{\mathbb{E}[d(A, B)] - \hat{d}(A, B)}{\mathbb{E}[d(A, B)]} \qquad (7)$$

The two language clusters A and B are considered to be related if the relative frequency of $\hat{d}(A, B) \geq d(A, B)$ compared to the number of random permutations, which considered as *p-value*, is less than 0.05. The given languages are termed *related* if the final two clusters that are merged at the root are related (Kessler and Lehtonen, 2006).

Kessler (2007) ran this test using various word similarity metrics which almost give similar results. Among these metrics, we ran on *P1-dolgo* which is a binary metric that determines whether the consonant class of the word's initial consonant matches or not. Additionally, we employ the binary similarity measure introduced by Turchin et al. (2010) to test the significance of the Altaic family where the first two consonants are considered. We further test continuous word distances introduced by List (2010) (*SCA*) and List (2012) (*LexStat*) that are based on sequence alignment techniques which were introduced in the context of automated cognate detection.

[1] https://www.wiktionary.org/

4.3 Implementation

We mapped the consonant classes to the protein alphabet since phylogenetic softwares expect input as either nucleotide or amino acid sequences. The multiple sequence alignments are obtained from CLUSTALW2 (Larkin et al., 2007) while the best trees and their corresponding likelihoods were computed using IQ-TREE (Nguyen et al., 2015). As described in §3.4 and §3.5, the proportions of invariant sites P^0_{inv} and P^a_{inv} are set to 0.01 and 0.06 respectively for null (H_0) and alternate (H_a) hypotheses. The parametric bootstrap replicates are generated using AliSim (Ly-Trong et al., 2022), an extension of IQ-TREE. To replicate as closely as possible, gaps present in the original character matrices are retained in the replicates. We calculate the p-value based on a sample size of $k = 15$. The outcomes are observed to be stable beyond this size. The word similarity metrics used in the baseline models are computed by using Lingpy (List and Forkel, 2021). For the phylogenetic tree construction task, MEGA11 (Tamura et al., 2021) was used to deduce ML-tree with the aforementioned model with an additional gamma rate heterogeneity parameter with two distinct rates whose shape is estimated. We name this method as *ML-Poisson+I+G2*. The generalized quartet distances (GQD) (Pompei et al., 2011) between the predicted and the gold trees are computed from quartet distances obtained using qdist (Mailund and Pedersen, 2004). Unlike biological phylogenetic trees, language trees are often multifurcated. Hence, GQD excludes penalties over the order of bifurcations. The code and relevant data is made available as part of supplementary material. Further implementation details can be found in `README.md` therein.

5 Results

The primary results of the paper are tabulated in Table 3, where the results of LRT (last row) are compared with those of the multilateral permutation test (others). Except for LRT, the column 'Method' indicates the distance metric employed in the permutation test. For permutation test, the values indicate the similarity metric $\hat{s}$ defined in Eq. 7, as measured at the root. On the other hand, for LRT the values indicate the mean of observed $\hat{\delta}$ (see §3.5). The p-values are indicated in parentheses. The standard threshold of 0.05 is assumed for p-values.

One can observe that false positives, indicated in red, are absent for LRT, in contrast with multilateral permutation tests which exhibit false positives wherever possible (except P1-Dolgo for MKh-UAz). However, we note that the similarity scores of the Turchin measure are consistently small (< 0.005) for negatives irrespective of the significance implied by the p-value. Hence, in this respect, it may be noted that Turchin could be a good measure when similarity scores are taken into consideration.

5.1 Tree Construction

As mentioned in §4.1, both the methods output a tree and, therefore, the methods have been evaluated on the tree construction task. The purpose of this task is to ensure that the proposed methods have indeed a good sense of phylogenetic inference and are, hence, appropriate to carry out significance tests over phylogenies. The results are tabulated in Table 4. By comparing with the mean scores of state-of-the-art language phylogeny inference methods on this data, ML-tree+I+G2 (0.079) is few steps behind Bayesian inferred tree (0.066) (Rama et al., 2018) and maximum a posteriori tree (0.051) (Rama and List, 2019). Hence, it can be concluded that consonant-class-based character matrix encoding is almost as good as cognate-based binary character matrix encoding while probabilistic methods based on character matrices are superior to distance based methods for this task. Among the distance-based approaches, one with the SCA metric performs best. A similar situation was observed in Rama et al. (2018) and Rama and List (2019) where SCA-based cognates yield the best performance. However, it should be noted that SCA and LexStat-based measures yield false positives on significance testing (Table 3) despite their performance on this task.

6 Evaluation of Macro Families

We apply the tests on groupings of a few families from proposed macro families, namely Nostratic, Macro-Mayan, and Amerind. Under Nostratic, we test for groupings Dravidian-Indo-European (*Drav-IE*) and Dravidian-Indo-European-Kartvelian (*Drav-IE-Kart*) while we test Mayan-Mixezoque under Macro-Mayan and Mayan-Utoaztecan under Amerind. The results are tabulated in Table 5. While the LRT test supports all of the mentioned families, the mean LRT statistic $\hat{\delta}$ is weak (negative or close to 0) for Drav-IE-Kart (Nostratic) and Mayan-Utoaztecan (Amerind).

Method	MKh	Mun	MKh-Mun	MKh-May	MKh-UAz	AfA-LoBur	IE	Drav	May	MZ	UAz
Related	✓	✓	✓	✗	✗	✗	✓	✓	✓	✓	✓
P1-Dolgo	0.123 (<0.001)	0.243 (<0.001)	0.080 (<0.001)	0.007 (<0.001)	0.005 (0.063)	0.017 (<0.001)	0.071 (<0.001)	0.440 (<0.001)	0.228 (<0.001)	0.412 (<0.001)	0.572 (<0.001)
Turchin	0.019 (<0.001)	0.124 (<0.001)	0.019 (<0.001)	0.003 (<0.001)	0.003 (0.005)	0.004 (<0.001)	0.028 (<0.001)	0.292 (<0.001)	0.126 (<0.001)	0.256 (<0.001)	0.402 (<0.001)
LexStat	0.065 (<0.01)	0.138 (<0.01)	0.048 (<0.01)	0.028 (<0.01)	0.018 (<0.01)	0.033 (<0.01)	0.036 (<0.01)	0.197 (<0.01)	0.129 (<0.01)	0.244 (<0.01)	0.306 (<0.01)
SCA	0.087 (<0.01)	0.187 (<0.01)	0.074 (<0.01)	0.015 (<0.01)	0.006 (<0.01)	0.031 (<0.01)	0.056 (<0.01)	0.296 (<0.01)	0.177 (<0.01)	0.304 (<0.01)	0.400 (<0.01)
LRT	9.205 (<0.001)	1.58 (<0.001)	14.18 (<0.001)	-14.359 (0.280)	-12.188 (0.065)	-10.768 (0.979)	26.154 (<0.001)	1.78 (<0.001)	68.212 (<0.001)	7.192 (<0.001)	10.448 (<0.001)

Table 3: Significance testing on various existent and non-existent families. The values indicate the similarity measure $\hat{s}$ in case of permutation tests and in case of LRT they indicate the mean of statistic $\hat{\delta}$. Values in parentheses indicate p-value. False positives are marked in red.

Method	AA	AN	IE	PN	ST	Avg
P1-Dolgo	0.060	0.208	0.033	0.175	0.188	0.133
Turchin	0.069	0.195	0.058	0.175	0.275	0.154
LexStat	0.051	0.178	**0.020**	0.164	0.096	0.102
SCA	0.049	0.119	0.025	0.166	**0.087**	0.089
ML-Poisson +I+G2	**0.026**	**0.065**	0.033	**0.145**	0.125	**0.079**

Table 4: Comparison of the methods on phylogenetic tree construction task provided as GQD scores. The best results are in **bold**.

Method	Drav-IE	Drav-IE-Kart	May-MZ	May-UAz
P1-Dolgo	0.046 (<0.001)	0.038 (<0.001)	0.033 (<0.001)	0.046 (<0.001)
Turchin	0.017 (<0.001)	0.002 (0.197)	0.012 (<0.001)	0.012 (<0.001)
LexStat	0.024 (<0.01)	0.014 (<0.01)	0.033 (<0.01)	0.027 (<0.01)
SCA	0.024 (<0.01)	0.007 (0.01)	0.019 (<0.01)	0.024 (<0.01)
LRT	24.882 (<0.001)	0.316 (<0.001)	20.988 (<0.001)	-1.035 (<0.001)

Table 5: Results of evaluation of macro families. Parentheses contain p-values.

In other words, by looking at Eq. 4, the alternate hypothesis H_a, i.e., having higher invariant sites is not preferred. It can be said that LRT is a highly sensitive test since the addition of a single language, viz., Georgian to strongly support Drav-IE with 16 languages alters the outcome drastically. This is a desirable property since the presence of even a single anomaly, an unrelated language in this case, can be detected. Note that other combinations in Nostratic such as Drav-Kart or IE-Kart are much weaker and not well supported by the permutation test itself, which is elaborated as follows.

6.1 Analysis of Permutation tests on Nostratic

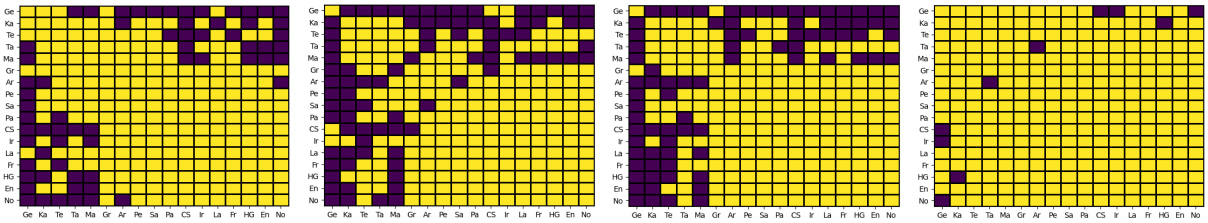
Bilateral significances on Nostratic grouping Drav-IE-Kart for various distance metrics are reported in Fig. 2, where the pairwise relationships based on p-value (with threshold 0.05) are color-coded. The computation follows the same steps as defined in §4.2 except that distances and similarities are calculated over pairs of languages instead of language clusters. This indeed forms the first iteration of a complete multilateral test.

The languages are abbreviated in Fig. 2 as follows: Old Georgian (Ge), Old Kannada (Ka), Old Telugu (Te), Old Tamil (Ta), Old Malayalam (Ma), Greek Ancient (Gr), Old Armenian (Ar), Middle Persian (Pe), Sanskrit (Sa), Pali (Pa), Old Church Slavonic (CS), Old Irish (Ir), Latin (La), Old French (Fr), Old High German (HG), Old English (En) and Old Norse (No).

It is visible that for each metric, languages of the same family (IE and Drav) are almost always related pairwise. Secondly, many pairs from Drav-IE appear related. However, except for LexStat, Georgian shows to be related to at most two languages from the Drav-IE grouping. Yet, in the permutation tests for these metrics, except for Turchin (Table 5), Drav-IE-Kart appears significantly related with sometimes even good similarity scores (in case of P1-Dolgo). All that can be concluded here is that, except for the LexStat metric, permutation tests are very much sensitive to pairwise language comparisons and may not yield false positives. However, if Drav-IE-Kart is to be considered a valid grouping, these tests may be said to yield false negatives.

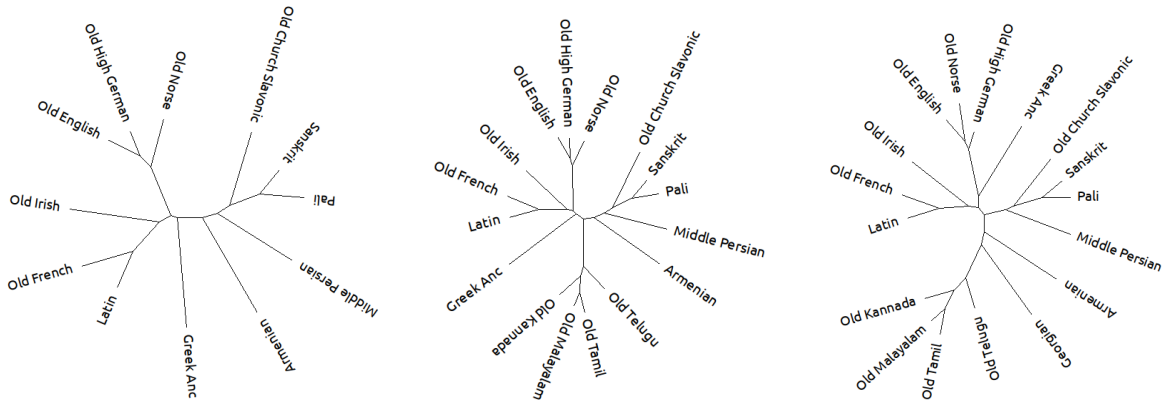
6.2 Analysis of ML-trees of Nostratic

Unrooted ML-trees are drawn in Fig. 3 on various sub-groupings of Nostratic using MEGA11 assuming the Poisson+I model. For the IE tree, the sub-families, except for the position of Old Church Slavonic, are highly faithful reflecting the existing notions. For instance, the topology of the Germanic family, i.e., [Old Norse, [Old English, Old



(a) P1-Dolgo (b) Turchin (c) SCA (d) LexStat

Figure 2: Bilateral (pairwise) significance among the languages of Nostratic grouping. The yellow shade implies that the relationship is statistically significant ($p < 0.05$), while the purple shade implies otherwise.



(a) IE (b) Drav-IE (c) Drav-IE-Kart

Figure 3: Comparison of unrooted ML-trees on various groupings of Nostratic language families

High German]] contains the valid West-Germanic branch [Old English, Old High German]. Similarly, the Italo-Celtic group [Old Irish, [Latin, Old French]] is visible. Also, one can distinguish a clear boundary between Western and Eastern IE languages reflecting the geographical distribution. However, the position of Old Church Slavonic intruded into Indo-Iranian appears problematic.

Further, the addition of the Dravidian family in Drav-IE does not alter the IE topology. It is intriguing to note the western inclination of Dravidian given its eastern geographical location in the present day. However, this is in line with the observation of Caldwell (1875), the founder of comparative Dravidian linguistics himself. Finally, the addition of Georgian invalidates the West-Germanic branch as well as pushes Greek problematically into the western group. However, much of the topology is undisturbed and one can also notice how the languages/families that are located south of the Caucasus namely, Armenian, Georgian, and Dravidian are grouped. Overall, it may be concluded that the addition of unrelated or weakly related languages can alter the actual topology.

7 Conclusions

In this paper, we have presented a likelihood ratio test based on the proportions of invariant sites to determine the genetic relatedness of a group of languages. Our proposed test does not yield false positives, which is in contrast with previous permutation-based tests that proved to be good only for pairwise language comparisons and not for validating a language group. By applying this test, we have found strong supporting evidence for macro-families such as Dravidian-Indo-European, Macro-Mayan (for Mayan-Mixe-Zoque), and weak evidence for Nostratic (Dravidian-Indo-European-Kartvelian) and Amerind (for Mayan-Uto-Aztecan). Through secondary analyses, we have also shown that probabilistic-based methods are superior to distance-based ones based on tree construction and correlation of topologies with geography. In summary, before the construction of phylogenies of a group of languages, the relatedness of the group must be established through a credible significance test such as the one we have presented. Otherwise the phylogeny would not only be meaningless but may also alter the topology of a related sub-group.

Limitations

The values of P^0_{inv} and P^a_{inv} (§3.5) are roughly decided based on the estimated ones from two examples, namely, Afrasian-Lolo-Burmese as a negative example and Indo-European as a positive example. The question that what should be the most appropriate values that should make the test optimal is not addressed here. Ideally, to address this question, more data is needed with several positive and negative examples in order to search for optimal values of these parameters. Also, the exact values may require calibration according to the phylogenetic software used since there could be significant differences in the implementations. Secondly, while analyzing Nostratic languages, Uralic, an important language family, has not been included due to the selection criteria (§4.1) that the languages should have been attested before 10th centuary C.E. To include Uralic, the (Nostratic) languages that are attested around the same period as the earliest attested ones from Uralic (roughly 1300 CE onwards) should be considered in order to make 'fair' comparisons.

Ethics Statement

All the datasets are obtained from publicly available sources. Thus, there are no foreseen ethical considerations nor conflicts of interests.

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