Intertemporal Preference Steering in Qwen3 via Contrastive Activation Addition

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Abstract

We study linear representations of temporal horizon in the large language model `Qwen3-32B` and use them to change the model's time-related preferences, recommendations, and capabilities. We train contrastive linear probes on teacher-forced temporal-choice answers to find a short-term versus long-term direction in the model's residual stream and evaluate contrastive activation addition steering on a held-out binary temporal-choice task, an out-of-distribution monetary intertemporal-choice task, and a TravelPlanner capability benchmark. The central result is that temporal-horizon directions can be identified with simple contrastive linear probes and then used for steering to induce large, bidirectional preference changes. On an out-of-distribution monetary choice task that varies reward size and delay, steering strongly shifts the model's indifference threshold between smaller-sooner and larger-later rewards in both directions. We further show improvements on a planning-related capability metric under moderate temporal steering. These results suggest that model intertemporal preferences are measurable and steerable, which is relevant for AI systems that give advice involving delayed costs and benefits, and for safety questions about long-horizon planning.

1. Temporal-Horizon Direction and Probe Diagnostics

1.1. Direction Construction

This section shows that intertemporal preference in the residual stream of the large language model `Qwen3-32B` is linearly separable and defines the temporal-horizon vector used in the downstream activation-steering experiments.

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Linear directions have been used to study high-level model representations (Marks & Tegmark, 2024), and contrastive activation addition uses such directions as steering vectors during inference (Rimsky et al., 2024). We therefore focus on a signed contrastive difference-in-means direction. The direction is trained on paired short-term and long-term answer continuations. Because these continuations might also differ in abstraction, urgency, or other semantic concepts, we treat probe accuracy as evidence that the training contrast is linearly accessible, not as evidence that the direction isolates temporal horizon alone. For diagnostic purposes, this direction can be evaluated as a linear probe. We report the long-term versus short-term classification accuracies and compare them with other standard linear readouts. We treat the classification accuracies as readout diagnostics but do not necessarily use the method that maximizes classification accuracy for steering. Causal claims are reserved for the intervention experiments in Sections 2, 3, and 4. This separation is standard in mechanistic interpretability: readout accuracy identifies accessible information, while activation patching or activation steering tests whether changing a representation changes behavior (Alain & Bengio, 2016; Tenney et al., 2019; Elhage et al., 2021; Meng et al., 2022; Turner et al., 2024; Rimsky et al., 2024; Li et al., 2023; Zou et al., 2023).

All experiments use `Qwen3-32B`, with thinking trace disabled (Yang et al., 2025; Qwen Team, 2025). Mean-mass (MM) difference directions and comparison probes are trained at layers

$$24, \ 28, \ 32, \ 36, \ 40, \ 44, \ 48,$$

which span the middle-to-late part of the 64-layer model. We use an explicit dataset of 500 temporal-choice questions with overt time markers and an implicit dataset of 300 questions that encode horizon through semantic contrasts without explicit time words. Both datasets are validated for clear horizon distinction and balanced presentation order. The explicit set is split at question level into 400 train and 100 test questions. The implicit set is used only for cross-domain evaluation on the full 300-question set. Each explicit question yields two teacher-forced examples, so the explicit training matrix contains 800 examples.

The implicit examples are intended to reduce reliance on explicit time words, but they do not fully isolate horizon: in

Regime	Prompt	Short-term continuation	Long-term continuation
Explicit	When developing a roadmap for this initiative, consider:	What we can deliver in the next 30 days	What position we want in 10 years
Implicit	When the system fails, the first priority becomes	containing the spreading damage before it escalates	redesigning the architecture to prevent recurrence

Table 1. Representative probe-training examples. Each row yields two teacher-forced continuations with opposite temporal-horizon labels.

practice, long-horizon answers often become more systemic, preventive, and abstract, while short-horizon answers are often more local, reactive, and operational.

For each training example, the model receives the natural-language question as context and the short-term and long-term continuations are teacher-forced as the two distinctive options. The MM difference direction uses mean-answer pooling, which averages hidden states across all continuation tokens. Mean-answer pooling is used for the MM difference direction as a design choice: we expected a span-level answer feature to align better with interventions applied during generation than a single final-token feature. Final-token pooling at the last continuation token is also evaluated as a diagnostic variant. At each selected layer, the MM probe direction is

$$d_{\text{MM}} = \mu_{\text{long}} - \mu_{\text{imm}},$$

where each mean is computed from the training activations at a fixed layer and feature. This contrastive construction follows the broader use of latent directions as simple behavioral readouts (Burns et al., 2023) and the CAA practice of adding contrastive residual-stream vectors during inference (Rimsky et al., 2024). It is especially relevant here because it preserves a signed short-term-to-long-term direction that can be added back into the residual stream at inference.

The same features are also used to fit comparison readouts. Logistic regression (LR) is fit directly. Whitened logistic regression (WLR) first fits a train-only covariance whitener, using regularized covariance

$$\Sigma_{\text{reg}} = \Sigma + 10^{-2}\bar{\sigma}^2 I,$$

where $\bar{\sigma}^2$ is the average coordinate variance, and then fits LR in the whitened space. The whitened MM (WMM) probe applies the same covariance adjustment to the difference-of-means direction. These variants contextualize linear separability, while the downstream interventions use the unwhitened MM probe direction as the vector for CAA steering.

1.2. Probe Diagnostics

All probe families nearly saturate explicit holdout accuracy, while cross-domain implicit performance is lower but still substantial. The MM row shows that the specific contrastive direction used for CAA steering remains a strong temporal horizon readout even though LR and whitened variants

attain slightly higher classification accuracy. Notably, the two whitened probe families have very similar classification accuracies across the reported splits. The steering experiments therefore use CAA steering with the explicit-trained mean-answer MM difference direction.

2. Binary Temporal Steering

This section shows that CAA steering with the explicit-trained MM difference direction causally affects temporal-choice behavior on held-out binary choices from the same broad short-term/long-term answer family. The evaluation uses similarly structured but unseen binary-choice questions. No evaluation question is used to fit the probe direction, and the only intervention is a signed residual-stream addition during inference. If positive and negative additions produce systematic changes in both parsed choices and forced-choice log probabilities, the direction is not only a classifier direction but a causal control direction for this task family. The intervention uses the explicit-trained mean-answer MM difference direction from Section 1.

2.1. Evaluation Prompt

Unlike probe training, steering evaluation presents the full binary-choice context, so the model must choose between two complete answer candidates rather than score an answer in isolation:

```
<question>
Options:
  <short-term or long-term option>
  <the other horizon option>
Answer:
```

The short-term and long-term options are presented in randomized order, and the evaluated datasets are the held-out explicit test set with 100 prompts and the full implicit set with 300 prompts.

2.2. Intervention

For a selected layer l, signed strength α, and normalized MM difference vector $\hat{d}_l$, the hook adds $\alpha\hat{d}_l$ to the residual stream at the last prompt token and the same update is also applied at each decode step. Positive strengths steer toward the long-term class and negative strengths steer toward the immediate class. Table 3 summarizes the intervention grid.

Probe	Explicit holdout	Implicit full
LR	99.0% (L28)	83.5% (L24)
WLR	99.0% (L32)	80.3% (L28)
MM	96.0% (L24)	77.0% (L32)
WMM	99.0% (L32)	80.3% (L28)

Table 2. Explicit-trained probe results for the main mean-answer pooling feature. Entries report the best accuracy and the layer attaining it. The MM row is the probe family from which the CAA steering vector is constructed.

Figure 1. Layerwise explicit-trained probe accuracies from the probe-variation experiment. The left panel reports explicit test set accuracy; the right panel reports full implicit set accuracy.

We do not claim a uniquely optimal layer. We treat layers 24–48 as an exploratory probe-layer band and report full sweeps where feasible. The main result is that CAA effects are broad across middle-to-late layers, not that any specific layer is intrinsically privileged. Single-layer plots are representative slices chosen for readability or compute constraints.

Because the steering vectors are normalized, the added residual perturbation has L2 norm $|\alpha|$. The mean hook-site residual L2 norms in baseline (unsteered) generations are between 202 at layer 24 and 355 at layer 48. The strongest intervention $|\alpha| = 128$ is therefore 36%–63% of the unsteered norm, which is a deliberately large perturbation, while $|\alpha| \le 16$ is less than 8% of the corresponding residual norm. For details on intervention and baseline residual stream magnitudes, see Appendix E.

2.3. Choice and Likelihood Metrics

Generated continuations are parsed into long-term or immediate choices, and the main behavioral metric is the proportion of parsed continuations choosing the long-term option. If the parser cannot identify the choice, the evaluator falls back to teacher-forced candidate logprob comparison; the plotted choice rates are computed from parsed, no-fallback

choices. The likelihood metric computes

$$\Delta_{\text{avg}} = \ell(c_{\text{long}} \mid p) - \ell(c_{\text{imm}} \mid p),$$

$$\ell(c \mid p) = \frac{1}{|c|} \sum_t \log P(c_t \mid p, c_{<t}),$$

and reports its change relative to the unsteered baseline. This average-token logprob margin controls for continuation length; the full layerwise logprob surface is shown in Appendix B, Figure 6.

We see the causal steering effect across all explored layers, with layer 44 giving the largest parsed-choice contrast on both evaluation sets. This broad middle-to-late layer profile is consistent with prior activation-steering and intervention work, which often finds the strongest causal directions away from the earliest and latest layers of the stack (Turner et al., 2024; Rimsky et al., 2024; Durmus et al., 2024). Note that these numbers only describe the measured intervention profile; they do not by themselves identify a complete circuit for intertemporal preference.

As a matched-norm control, we also steer with a random unit vector orthogonal to the MM difference direction at each layer. This control produces substantially smaller choice shifts than CAA steering, and those shifts are not consistently aligned with the sign of the intervention (Appendix C, Figure 7). The main binary results use greedy decoding with temperature 0, so repeated decoding is not a source of

Quantity	Values
Steered layers l	24, 28, 32, 36, 40, 44, 48
Signed strengths α	0, ± 2, ± 4, ± 8, ± 16, ± 32, ± 64, ± 128
Patch positions	Last prompt token and every generated-token step

Table 3. Binary steering configuration. Each intervention adds $\alpha \hat{d}_l$ to the residual stream at layer l.

Figure 2. Proportion of prompts choosing the long-term option over layer and signed steering strength.

sampling variance for those point estimates. As a robustness check, however, Appendix D repeats a stochastic layer sweep with temperature 0.8, top-$p = 1$, and five sampled generations per prompt. The sampled means preserve the greedy ordering and magnitudes, with 2-sigma standard-error bars small relative to the steering effect.

3. Time-Utility Steering

This section shows generalization beyond the binary temporal-choice distribution toward settings where models might advise users about real-world tradeoffs. The model is asked monetary time-utility questions involving dollar amounts and delays, a quantitatively measurable task that was not used to train the temporal horizon direction. The goal is to show that a direction trained from semantic short-term versus long-term answers can change preferences on a structurally distinct advice task, not merely on held-out examples with the same format. This setting is related to classical time preference and discounting measurements (Samuelson, 1937; Frederick et al., 2002).

3.1. Task and Steering Setup

The prompt is written from the user's point of view:

```
I have an offer to take <x>$ now
or <y>$ in <t>.
What should I choose?
Answer in just a few words.
```

The monetary grid is defined by a base immediate amount x_b, base delayed amounts y_b, and amount scale factors c. Each prompt uses $x = cx_b$ and $y = cy_b$, so the delayed/immediate multiplier $m_b = y_b/x_b$ is held fixed while the absolute dollar amounts vary. This separates numeric proportions from absolute-dollar lexical effects. The experiment also uses day-count delay labels where possible rather than mixed natural labels. For example, "in 365 days" and "in a year" denote approximately the same delay, but the latter introduces additional calendar semantics.

Each condition therefore has 8 amount multipliers, 12 delays, and 8 scale factors, for 768 generations. Parsed responses are categorized as choosing the immediate amount, choosing the delayed amount or "it depends".

3.2. Area-Equivalent Indifference Multiplier

The area-equivalent indifference multiplier estimates the delayed/immediate reward ratio y/x at which the model is indifferent between taking x now and y after delay t. For a fixed condition and delay, all scale-factor rows with the same base multiplier $m = y/x$ are combined. Let $q_{\text{now}}(m, t)$ be the parsed present-choice fraction at multiplier m and delay t. "It depends" responses are split equally between present and delayed choices, although this convention has negligible effect because "it depends" occurs in only 0.15% of responses. Write $z_i = \log m_i$ for the sorted sampled multipliers in a delay column and define midpoint bin edges on the log axis by $e_0 = z_{\min}$, $e_n = z_{\max}$, and $e_i = \frac{1}{2}(z_i +$

Quantity	Values
Base immediate amount x_b	100
Base delayed amounts y_b	100, 101, 110, 150, 200, 500, 1000, 10000
Multipliers $m_b = y_b/x_b$	1.0, 1.01, 1.1, 1.5, 2, 5, 10, 100
Amount scale factors c	0.3, 0.5, 0.7, 1, 1.4, 2, 6, 8
Delays t	1 hour, 8 hours, 1 day, 2 days, 7 days, 30 days, 90 days, 365 days, 730 days, 1095 days, 1825 days, 3650 days
Signed strengths α	0, ± 8, ± 16, ± 32, ± 64, ± 128

Table 4. Time-utility steering grid. For each y_b, scale factor c, and delay t, the prompt compares $x = cx_b$ now against $y = cy_b$ after delay t.

z_{i+1}) for $1 \leq i < n$. Each sampled multiplier is then assigned bin width $w_i = e_i - e_{i-1}$. Treating $q_{\text{now}}(z, t)$ as piecewise constant on these bins, the ideal full-coverage log-axis average is

$$A_{\text{now}}(t) = \frac{\sum_i w_i q_{\text{now}}(m_i, t)}{\sum_i w_i},$$

since $\sum_i w_i = z_{\max} - z_{\min}$, the log indifference multiplier $z^*(t)$ is

$$z^*(t) = z_{\min} + A_{\text{now}}(t)(z_{\max} - z_{\min})$$
$$= z_{\min} + \sum_i w_i q_{\text{now}}(m_i, t).$$

This is a piecewise-constant log-axis average, chosen because the tested multipliers are sparse and unevenly spaced and the empirical choice columns can be locally non-monotone. The corresponding indifference multiplier on the original scale is $m^*(t)$:

$$m^*(t) = \exp(z^*(t)).$$

Large $m^*(t)$ means that the model requires a larger delayed reward before it switches away from the immediate option, while small $m^*(t)$ means that a smaller delayed reward is sufficient to make the model wait.

Appendix F shows the full present-choice heatmaps and precise AEIM values. The heatmaps show the same directional pattern summarized by AEIM: negative steering expands the present-choice region, while positive steering contracts it, without relying on the AEIM aggregation.

3.3. Layer Selection and Results

We sweep the same seven probe layers as in the binary-steering experiment. Figure 3 summarizes the geometric mean AEIM over the 12 delays for each layer and signed strength. The effect is signed and broad across layers. The layer-40 curves below are therefore an illustrative slice of the sweep, not the unique selected optimum. The other layerwise AEIM curves are shown in Appendix F.

The AEIM curves show that the result is delay-dependent rather than uniform: as expected, negative steering moves the indifference multiplier upward, positive steering moves it downward, and most displacement occurs at long delays. At ± 128, the median ratio of negative-steer AEIM to positive-steer AEIM across delays is 4.90 and at 3650 days it is over 56. In other words, at the 10-year horizon, negative steering makes the model require more than 56 times the delayed reward needed under positive steering. The result is a strong generalization test: a direction trained only from binary temporal-horizon answers changes the model's measured intertemporal preferences on a new monetary-choice task.

For robustness, we performed a matched random-orthogonal control, which produces smaller, inconsistently signed movement (Appendix F). Unlike the MM difference vector, the random vector does not produce a consistent signed short-term versus long-term pattern across layers.

The main sweep uses greedy decoding with temperature 0. Analogously to Section 2, Appendix G reports robustness checks with stochastic sampling, which preserve the same qualitative pattern. Appendix E reports the steering vector magnitudes relative to unsteered hook-site residual-stream magnitudes, which are always less than 70% even at the strongest intervention.

4. TravelPlanner Capability

This section examines how temporal-horizon steering affects a planning-relevant capability metric, rather than only changing expressed preferences. We evaluate the same layer-40 temporal direction on TravelPlanner, a real-world travel planning benchmark that requires models to produce multi-day itineraries satisfying transportation, accommodation, restaurant, attraction, budget, and constraint requirements (Xie et al., 2024). A planning benchmark is a natural test for long-horizon planning because it requires maintaining a temporally extended plan while satisfying many local constraints. We use the validation split: 180 queries, comprising 20 queries for each combination of three difficulty levels and three trip lengths. Each query asks for a complete plan containing transportation, accommodation, meals, and attractions for each day. We use layer 40 as a compute-

Figure 3. Time-utility layer sweep. Each cell reports the geometric mean AEIM.

Residual addition	Geom. mean AEIM	Parsed later rate	It-depends rate	Unparsed rate
$-128\hat{d}_{40}$	8.34	36.3%	0.0%	0.26%
0	2.67	54.7%	0.0%	0.00%
$+128\hat{d}_{40}$	1.12	80.4%	1.2%	0.52%

Table 5. Layer-40 aggregate time-utility results over all delay and amount cells. The geometric mean AEIM is computed across the 12 delay columns. Parsed later and it-depends rates are computed over parsed generations in the condition.

constrained representative layer from the middle-to-late probe band rather than as a task-specific optimum.

4.1. Metric Choice

The benchmark has multiple metrics; we use Commonsense Constraint Micro Pass Rate as the primary metric for this experiment. It is the total number of passed commonsense checks divided by the total number of commonsense checks across all evaluated queries. Commonsense micro pass rate is the most informative metric for our purposes. In the baseline condition (no steering), delivery rate is already near ceiling (96.7%), while commonsense macro pass rate, hard-constraint rates, and final pass rate are near floor. Commonsense micro pass rate therefore captures partial improvements in itinerary plausibility even when strict whole-query metrics remain too coarse to distinguish conditions.

Figure 5 shows a non-monotonic but directionally meaningful effect. Moderate long-term steering improves com-

monsense micro pass rate, while negative steering degrades it. Very large positive steering also degrades performance. A likely explanation is that +128 pushes the model far enough off-distribution that generation coherence begins to deteriorate. The result is not evidence that temporal steering generally solves planning, and we do not know from this aggregate alone which commonsense checks flip under steering, but it suggests that the same direction that shifts intertemporal preferences can also affect a planning-relevant capability metric, with useful effects concentrated at moderate positive strengths.

5. Discussion

5.1. Summary

These experiments suggest that temporal horizon can be represented by simple linear directions in Qwen3 and that CAA steering with a direction trained from contrastive short-term versus long-term answers can produce large behavioral

Figure 4. Area-equivalent indifference multipliers from the layer-40 time-utility experiment. Note the -128 steering at 3650 days is capped at 100 because all model responses chose the immediate option.

Figure 5. TravelPlanner Commonsense Constraint Micro Pass Rate as a function of signed temporal steering strength. Positive values steer toward the long-term direction; negative values steer toward the short-term direction.

effects. The direction should not be interpreted as an isolated temporal-horizon feature, since the training contrast may also contain correlated semantic features. Nevertheless, the time-utility experiment suggests that intertemporal preference is at least a relevant component of the direction. The probe results establish that short-term and long-term answers are linearly separable, and the binary steering experiment shows that steering along the direction causally affects held-out temporal-choice questions. The time-utility experiment is the strongest generalization result. A direction trained only on binary temporal-horizon answers shifts monetary intertemporal choice on a structurally different task, moving AEIM by large factors in both directions. The TravelPlanner experiment further suggests that temporal-horizon steering can affect a planning-relevant capability metric, although the effect is moderate and non-monotonic.

The main takeaway is that MM difference directions built from contrastive temporal-horizon answers can generalize beyond their training format when used for CAA steering.

This matters because many deployed model outputs implicitly encode intertemporal preferences, such as advice about planning, finance, health, education, and policy that trades near-term actions against delayed outcomes. The results suggest that such intertemporal preferences can be measured and steered, rather than treated as an opaque side effect of pretraining or instruction tuning. This is also safety-relevant because temporal preferences may be one ingredient in long-horizon planning and reward-seeking, and a model that internally optimizes over a different horizon than the one it presents to users may be difficult to evaluate from outputs alone. At the same time, the capability to steer temporal preferences could be abused by a malicious actor to push a model toward decisions that are misaligned with user values, so the results underscore the importance of secure access controls and monitoring for similar behavior.

5.2. Limitations

The results rely on several assumptions that may fail in practice:

Construct validity. The MM direction is trained on short-term versus long-term answer continuations, but those continuations can also vary in abstraction, urgency, strategic scope, and proactive-versus-reactive framing. If these correlated features dominate the direction, the intervention should be interpreted as steering a broader short-term/long-term answer register rather than an isolated temporal-horizon variable. The monetary time-utility transfer suggests intertemporal preference is involved, but it does not fully identify the direction's semantic basis.

Intervention scale and specificity. Large activation additions may change more than temporal preference. The residual-norm diagnostics calibrate intervention size at the hook sites, but they do not prove that unrelated capabilities are preserved, especially at the strongest steering values. In deployment, an intervention that shifts time preference while also degrading factuality, instruction following, or coherence would be less useful and potentially unsafe.

Scope of empirical claims. All experiments use one model family and size, a small set of prompt templates, and thinking-disabled decoding. The results may not transfer to other architectures, smaller or larger models, reasoning-enabled settings, or tasks where temporal tradeoffs are expressed less explicitly. Performance may also depend on prompt wording or answer-format instructions.

Layer, strength, and prompt sensitivity. Layer and strength choices are exploratory. Full sweeps are reported where feasible, but some downstream experiments use representative layers for compute reasons. Performance may depend on layer choice, steering norm, and decoding settings.

Prompting baselines. The paper studies activation steering, not the full space of ways to alter model behavior. Similar behavioral changes might be possible by prompting the model to think more short-term or long-term, and such baselines are not reported here.

Parsing imprecision. Some generated answers are unparsed, especially in the free-generation steering experiments. The unparsed fraction is small enough that it is unlikely to change the qualitative results, but it can still introduce imprecision into parsed choice rates and derived quantities.

AEIM aggregation. AEIM compresses a sparse, unevenly sampled amount-delay grid into one threshold-like curve per condition. Different multiplier grids, delay grids, or aggregation conventions could change the apparent curve shape. The raw present-choice heatmaps in Appendix F

show the same steering pattern without this aggregation, but they have a less clean scalar interpretation.

Probe-family comparisons. Steering is evaluated primarily with the MM difference direction. Matched-norm comparisons against LR, WLR, and WMM steering directions would clarify whether MM is uniquely useful for intervention.

Benchmark coverage. The TravelPlanner evidence is weaker than the time-utility result because the aggregate commonsense metric is indirect and does not identify which item-level checks change under steering.

5.3. Compute Resources

All experiments used Qwen3-32B inference. Activation collection and steering runs were executed on NVIDIA GPU hardware with 100 GB of available VRAM on RTX 5090 or similar-capability hardware. The layer-sweep steering experiments took approximately 48 hours for binary-choice steering and 24 hours for time-utility steering. The respective stochastic steering sweeps took about twice as long: 4 days and 2 days. TravelPlanner evaluation with single-layer steering took approximately 48 hours. The remaining diagnostics and plotting took seconds to minutes on a MacBook Air M4.

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Appendix

A. Binary Steering Layerwise Summary

Table 6 gives the layerwise contrast table discussed in Section 2. Layer 48 is marginally largest by logprob gap on both datasets, while layer 44 is largest by choice gap.

B. Binary Steering Logprob Diagnostics

C. Binary Random-Orthogonal Control

D. Stochastic Binary Steering Robustness

The main binary steering sweep is decoded greedily. To estimate sensitivity to sampling randomness, Figures 8–10 show the stochastic layer sweep with temperature 0.8, top-$p = 1$, and five sampled generations per prompt. In Figures 8–10 and Table 7, error bars are 2-sigma standard errors, computed as $2\sigma/\sqrt{5}$ across the five repeat-level summaries for each cell. The sweep covers all seven probe layers but only the signed strengths $0, \pm 8, \pm 32, \pm 128$ due to compute constraints. The sampled subset preserves the qualitative steering pattern, suggesting the greedy result is not an artifact of deterministic decoding. Table 7 reports the layer-44 explicit-test slice corresponding to the main binary steering contrast.

E. Residual-Norm Diagnostics

F. Time-Utility Diagnostics

Figure 15 shows the full grid underlying the AEIM summaries. Yellow cells are high present-choice proportions, and dark cells are low present-choice proportions. Because delay increases from left to right and delayed/present amount increases from bottom to top, a rationally expected pattern has a yellow cluster in the bottom-right region: when the delayed amount is not much larger than the present amount and the wait is long, the present option should usually be preferred. Conversely, the top-left region should be dark, because a much larger delayed reward after a short wait should usually be preferred. The steering pattern follows the intended direction: negative strengths expand the yellow present-choice region, while positive strengths contract it; the expansion and contraction become larger as steering strength increases.

G. Stochastic Time-Utility Robustness

The main time-utility sweep is decoded greedily. To estimate sensitivity to sampling randomness, Figures 19 and 20 repeat a stochastic layer sweep with temperature 0.8, top-$p = 1$, and five sampled generations per prompt. The sweep covers all seven probe layers at signed strengths $0, \pm 8, \pm 32, \pm 128$, a strength subset of the full greedy grid in the main text due to compute constraints. This stochastic sweep preserves the qualitative steering pattern, suggesting the greedy result is not an artifact of deterministic decoding. In Figures 19–20 and Table 8, error bars are 2-sigma standard errors, computed as $2\sigma/\sqrt{5}$ across the five repeat-level summaries for each cell.

Layer	Explicit choice gap	Explicit logprob gap	Implicit choice gap	Implicit logprob gap
24	23.9 pp	0.158	4.6 pp	0.112
28	25.3 pp	0.125	10.2 pp	0.121
32	37.6 pp	0.175	11.0 pp	0.101
36	37.5 pp	0.183	14.6 pp	0.084
40	28.0 pp	0.399	10.0 pp	0.146
44	59.4 pp	0.813	36.5 pp	0.263
48	25.7 pp	0.824	14.6 pp	0.268

Table 6. Layerwise steering contrast. The choice gap is the best positive long-term no-fallback rate minus the best negative long-term no-fallback rate at that layer. The logprob gap is the best positive minus best negative long-minus-immediate average-logprob delta at that layer.

Figure 6. Change in long-minus-immediate average logprob margin relative to the unsteered baseline.

Residual addition	Greedy no-fallback long	Sampled no-fallback long	Sampled fallback
$-128\hat{d}_{44}$	33.7%	$40.6 \pm 1.9\%$	$2.4 \pm 0.8\%$
0	76.8%	$74.3 \pm 1.8\%$	$6.6 \pm 2.2\%$
$+128\hat{d}_{44}$	93.1%	$91.4 \pm 1.8\%$	$12.2 \pm 1.5\%$

Table 7. Stochastic binary steering repeat check on the explicit test set. The sampled columns report mean $\pm$ 2-sigma standard-error bars across five repeat-level sampled decodings.

Layer	$-128\hat{d}_l$	0	$+128\hat{d}_l$
24	8.48 ± 0.17	3.22 ± 0.22	1.63 ± 0.07
28	10.95 ± 0.72	3.22 ± 0.22	1.46 ± 0.04
32	15.20 ± 0.94	3.22 ± 0.22	1.10 ± 0.01
36	12.08 ± 0.49	3.22 ± 0.22	1.12 ± 0.01
40	7.84 ± 0.34	3.22 ± 0.22	1.18 ± 0.02
44	10.08 ± 0.66	3.22 ± 0.22	2.36 ± 0.09
48	15.11 ± 1.02	3.22 ± 0.22	2.09 ± 0.10

Table 8. Stochastic time-utility geometric-mean AEIM by layer. Entries report mean $\pm$ 2-sigma standard-error bars across five repeat-level sampled decodings.

Figure 7. Matched random-orthogonal binary steering control. The control vector at each layer is orthogonalized against the corresponding normalized MM difference vector. The largest absolute no-fallback choice movement is 5.6 percentage points on explicit test and 5.5 points on implicit full, compared with 59.4 and 36.5 percentage-point CAA contrasts for layer-44 steering on the same datasets.

Figure 8. Stochastic binary steering no-fallback choice results. Top: mean proportion choosing the long-term option. Bottom: 2-sigma standard-error bar across five sampled decodings.

Figure 9. Stochastic binary steering forced-choice logprob results. Top: mean change in long-minus-immediate average logprob margin relative to baseline. Bottom: 2-sigma standard-error bar across five sampled decodings.

Figure 10. Stochastic binary steering fallback diagnostics. Top: mean fallback rate. Bottom: 2-sigma standard-error bar across five sampled decodings.

Figure 11. Section 2 steering experiment hook-site residual-norm diagnostics from unsteered baseline generations. The plot reports mean residual-stream L2 norm at the prompt/decode positions where steering is applied.

Figure 12. Section 2 steering experiment intervention scale. The plot reports steering strength as a percentage of the mean hook-site residual norm from unsteered baseline generations.

Figure 13. Section 3 steering experiment hook-site residual-norm diagnostics from unsteered baseline generations. The plot reports mean residual-stream L2 norm at the prompt/decode positions where steering is applied.

Figure 14. Section 3 steering experiment intervention scale. The plot reports steering strength as a percentage of the mean hook-site residual norm from unsteered baseline generations.

Figure 15. Average present-choice heatmaps for the layer-40 time-utility experiment. Each panel averages over amount scale factors. Columns are delays, rows are delayed/present amount multipliers, and color is the parsed fraction choosing the present amount.

Figure 16. Layer-40 area-equivalent indifference multipliers for all delays and steering strengths. Higher values mean a larger delayed amount is needed to offset the delay. The -128, 3650-day point is right-censored at the largest tested multiplier because all valid parsed rows choose the immediate option; the true value is therefore ≥ 100.

Figure 17. Layerwise time-utility AEIM curves under CAA steering with the MM difference direction. These are the per-layer curves summarized by Figure 3.

Figure 18. Time-utility random-orthogonal control. The control uses matched-norm unit vectors orthogonal to the MM difference vector at each layer. It produces some AEIM movement, but the direction is not consistently aligned with steering sign: at ±128, the negative-over-positive geometric-mean AEIM ratio ranges from 0.39 to 1.69 across layers, compared with 4.55 to 14.46 for CAA steering with the MM difference direction.

Figure 19. Stochastic time-utility geometric-mean AEIM layer sweep. Top: mean geometric-mean AEIM across delays. Bottom: 2-sigma standard-error bar across five sampled decodings.

Figure 20. Stochastic time-utility AEIM curves for all seven probe layers. Bands show 2-sigma standard-error bars across five sampled decodings per prompt at temperature 0.8 and top-$p = 1$.