

000 001 002 003 004 005 006 UNIFIED CONVERGENCE ANALYSIS FOR SCORE- 007 BASED DIFFUSION MODELS WITH DETERMINISTIC 008 SAMPLERS 009 010 011

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ABSTRACT

Score-based diffusion models have emerged as powerful techniques for generating samples from high-dimensional data distributions. These models involve a two-phase process: first, injecting noise to transform the data distribution into a known prior distribution, and second, sampling to recover the original data distribution from noise. Among the various sampling methods, deterministic samplers stand out for their enhanced efficiency. However, analyzing these deterministic samplers presents unique challenges, as they preclude the use of established techniques such as Girsanov’s theorem, which are only applicable to stochastic samplers. Furthermore, existing analysis for deterministic samplers usually focuses on specific examples, lacking a generalized approach for general forward processes and various deterministic samplers. Our paper addresses these limitations by introducing a unified convergence analysis framework. To demonstrate the power of our framework, we analyze the variance-preserving (VP) forward process with the exponential integrator (EI) scheme, achieving iteration complexity of $\tilde{O}(d^2/\epsilon)$. Additionally, we provide a detailed analysis of Denoising Diffusion Implicit Models (DDIM)-type samplers, which have been underexplored in previous research, achieving polynomial iteration complexity.

1 INTRODUCTION

Score-based diffusion models (Sohl-Dickstein et al., 2015; Ho et al., 2020; Song et al., 2020a;c) have emerged as a powerful class of generative models, achieving significant success in image generation tasks, such as DALL-E (Ramesh et al., 2022), Stable Diffusion (Rombach et al., 2022), Imagen (Saharia et al., 2022), and SDXL (Podell et al., 2023). Beyond that, these models have also demonstrated effectiveness in diverse applications, including structure-based drug design (Corso et al., 2022; Guan et al., 2024), text generation (Austin et al., 2021; Zheng et al., 2023; Chen et al., 2023c), and reinforcement learning (Wang et al., 2022; Lu et al., 2023). At the core of the score-based diffusion models is a forward Stochastic Differential Equation (SDE) to diffuse the data distribution into a known prior distribution, and a neural network is trained to approximate the score function. Scalable score-matching techniques, such as denoising score matching (Vincent, 2011) and sliced score matching (Song et al., 2020b), enable efficient learning of the score function. Once learned, we can use numerical samplers to simulate the backward process and recover the original data distribution from noise.

To improve the sampling quality and efficiency of diffusion models, it is crucial to use efficient samplers in addition to an accurate score estimator. Early developments of diffusion models, such as Denoising Diffusion Probabilistic Models (DDPM) (Ho et al., 2020) applied stochastic samplers. Later, empirical studies showed that diffusion models with deterministic samplers, such as Denoising Diffusion Implicit Models (DDIM) (Song et al., 2020a) can still generate high-quality samples while achieving better efficiency. For instance, DDIM is more than 10 times faster than DDPM. Beyond DDIM, many novel fast ODE solvers have been developed for diffusion models, further improving the efficiency of the sampling processes (Lu et al., 2022; Zhou et al., 2024).

The remarkable success of diffusion models has inspired extensive research interest in the mathematical analysis of these powerful generative models (Block et al., 2020; De Bortoli et al., 2021; De Bortoli, 2022; Lee et al., 2023; Pidstrigach, 2022; Chen et al., 2022; 2023a;b; Li et al., 2024a;b;c; Benton et al., 2024; Huang et al., 2024a). Notably, many prior works studied the convergence of diffusion models with stochastic samplers (Chen et al., 2022; 2023a; Benton et al., 2024; Li et al.,

054 2024b). A key tool in their analysis is Girsanov’s theorem, which helps control the distance be-
 055 tween the distributions of two stochastic processes. However, it relies on the smoothing effect pro-
 056 vided by stochasticity and therefore do not apply to deterministic samplers. To address this, (Chen
 057 et al., 2024) introduced an additional corrector into the Langevin dynamics, which incorporates
 058 randomness to help smooth the distribution. Meanwhile, Chen et al. (2023b) assumed the access
 059 to the exact score function and provided a discretization analysis for the probability flow ODE in
 060 Kullback–Leibler (KL) divergence. Moreover, under extra conditions, Li et al. (2024b) proved a
 061 non-asymptotic convergence rate for a specific deterministic sampler based on the probability flow
 062 ODE with elementary analysis. Huang et al. (2024a) considered the Ornstein–Uhlenbeck (OU) for-
 063 ward process with the Runge–Kutta integrator and provided error bounds for both continuous- and
 064 discrete-time settings. However, all of these works focus on specific forward processes and sam-
 065 plers. Therefore, a natural question arises:
 066 *Can we develop a unified framework for the convergence analysis of diffusion models with deter-
 067 ministic samplers that accommodates those common forward processes and sampling algorithms?*

In this paper, we provide an affirmative answer to this question. We summarize our contributions as follows:

- We develop a key technical tool (Lemma 3.2), which enables us to bound the time derivative of the total variation (TV) distance between the final states of two ODE processes, through the difference in their drift terms and the divergence. As a direct application, we establish convergence guarantees for the continuous-time reverse ODE in the case of the OU forward process.
- We provide a unified convergence analysis framework for diffusion models with deterministic samplers. For general forward processes and sampling algorithms, our framework decomposes the error of diffusion models into five distinct terms: two arising from score estimation and three from time discretization. This decomposition enables a divide-and-conquer approach, allowing for improved analysis of each error component while maintaining the consistency of the unified framework.
- To demonstrate the generality and effectiveness of our framework, we apply it to two typical diffusion model settings: we achieve $\tilde{O}(d^2/\epsilon)$ iteration complexity for the Variance Preserving (VP) forward process with exponential integrator (EI) numerical scheme. This theoretical guarantee matches Li et al. (2024b), where a similar but different sampling algorithm is considered. Moreover, we establish polynomial iteration complexity for the Variance Exploding (VE) forward process with the DDIM numerical scheme. To the best of our knowledge, this is the first convergence result for diffusion models employing DDIM-type samplers that can handle estimated scores, whereas prior works, such as Chen et al. (2023b), only considered cases with access to accurate score function.

Notation: In this work, we use lowercase letters a, b to represent scalars, lowercase bold letters \mathbf{x}, \mathbf{y} to represent vectors, uppercase italic bold letters \mathbf{X}, \mathbf{Y} to represent random variables, and uppercase bold letters \mathbf{A}, \mathbf{B} to represent matrices. For a vector $\mathbf{x} \in \mathbb{R}^d$ and matrix $\mathbf{A} \in \mathbb{R}^{d \times d}$, we denote by $\|\mathbf{x}\|$ the Euclidean norm of \mathbf{x} and $\|\mathbf{A}\|_2$ the operator norm of \mathbf{A} . We use $f_1 \lesssim f_2$ to denote that there is a universal constant C such that $f_1 \leq Cf_2$. For two sequences $\{a_n\}$ and $\{b_n\}$, we write $a_n = O(b_n)$ if there exists an absolute constant C such that $a_n \leq Cb_n$, and we use $\tilde{O}(\cdot)$ to further hide the logarithmic factors. For vector operations, we use ∇ to denote the gradient, $\nabla \cdot$ to denote divergence, and ∇^2 to denote the Jacobian matrix.

2 PROBLEM BACKGROUND

The primary objective of diffusion models is to generate new samples given a set of examples drawn from the data distribution. In this section, we introduce the fundamentals of ODE-based diffusion models with deterministic samplers.

2.1 ODE-BASED DIFFUSION MODELS

A diffusion model typically consists of a forward process that perturbs the data distribution into noise, followed by a denoising backward process. In general, the forward process can be modeled as an Itô SDE:

$$d\mathbf{X}_t = \mathbf{f}(t, \mathbf{X}_t)dt + g(t)d\mathbf{W}_t, \quad \mathbf{X}_0 \sim q_0, \quad (2.1)$$

where \mathbf{W}_t is a Brownian motion in \mathbb{R}^d , $\mathbf{f}(t, \cdot)$ is called the drift coefficient, $g(t)$ is called the diffusion coefficient (Song et al., 2020c). It begins by sampling \mathbf{X}_0 from the data distribution q_0 , and

108 evolves according to the forward process (2.1). The law of \mathbf{X}_t is denoted by $q_t(\mathbf{x})$. Under mild reg-
 109 ularity conditions on q_0 , we can construct a family of reverse processes $(\mathbf{Y}_t^\lambda)_{t \in [0, T]}$, which evolve
 110 according to the following SDE:
 111

$$112 \quad d\mathbf{Y}_t^\lambda = -\left(\mathbf{f}(T-t, \mathbf{Y}_t^\lambda) - \frac{1+\lambda^2}{2}g(T-t)^2\nabla \log q_{T-t}(\mathbf{Y}_t^\lambda)\right)dt + \lambda g(T-t)d\mathbf{W}_t, \quad \mathbf{Y}_0^\lambda \sim q_T. \quad (2.2)$$

115 These processes hold the same marginal distribution as $q_{T-t}(\mathbf{x})$ at time t (Chen et al., 2023b). As a
 116 special case when $\lambda = 0$, the backward process is deterministic, i.e.,
 117

$$118 \quad d\mathbf{Y}_t = -\left(\mathbf{f}(T-t, \mathbf{Y}_t) - \frac{1}{2}g(T-t)^2\nabla \log q_{T-t}(\mathbf{Y}_t)\right)dt, \quad \mathbf{Y}_0 \sim q_T, \quad (2.3)$$

120 which is called the probability flow ODE (Song et al., 2020c).

121 The goal of ODE-based diffusion model is to simulate (2.3). However, we must make several ap-
 122 proximations. First, as we do not have access to the true score $\nabla \log q(t, \mathbf{x})$, we learn an approxi-
 123 mation $\mathbf{s}_\theta(t, \mathbf{x})$ by minimizing

$$124 \quad \mathcal{L}(\mathbf{s}_\theta) = \int_0^T \mathbb{E}_{q_t} [\|\mathbf{s}_\theta(t, \mathbf{X}_t) - \nabla \log q_t(\mathbf{X}_t)\|^2] dt. \quad (2.4)$$

127 When $\mathbf{f}(t, \cdot)$ is affine, the transition is Gaussian, allowing for the closed-form solution of
 128 $\nabla \log q(t, \mathbf{x})$. By applying integration by parts, (2.4) can be reformulated into tractable objec-
 129 tives, specifically through the methods of denoising score matching and implicit score matching
 130 (Hyvärinen & Dayan, 2005; Vincent, 2011). For more general forward processes, these objectives
 131 can be estimated using samples drawn from the forward process. Then, using the estimated score
 132 function \mathbf{s}_θ , we have the simulate reverse process:

$$133 \quad d\widehat{\mathbf{Y}}_t = -\left(\mathbf{f}(T-t, \widehat{\mathbf{Y}}_t) - \frac{1}{2}g(T-t)^2\mathbf{s}_\theta(T-t, \widehat{\mathbf{Y}}_t)\right)dt, \quad \widehat{\mathbf{Y}}_0 \sim q_T. \quad (2.5)$$

136 Second, since the initial distribution q_T of the reverse process is not directly accessible, we instead
 137 initialize the reverse process with a given Gaussian distribution π_d , which is assumed to be a good
 138 approximation of q_T .

139 Third, in practical implementations, the continuous-time process is typically simulated using time
 140 discretization. Specifically, we choose time steps $0 = t_0 < t_1 < \dots < t_N \leq T$, sample $\widehat{\mathbf{Y}}_0 \sim \pi_d$
 141 and iteratively calculate $\widehat{\mathbf{Y}}_{t_{k+1}}$ from $\widehat{\mathbf{Y}}_{t_k}$ using a particular numerical scheme. In this way, we only
 142 need to have access to the value of the estimated score function $\mathbf{s}_\theta(t_k, \widehat{\mathbf{Y}}_{t_k})$, $k = 0, 1, \dots, N$, which
 143 can reduce the computational complexity of the sampling process. Common numerical schemes in-
 144 clude the Exponential Integrator scheme, the Euler-Maruyama scheme, and the DDIM-type sampler
 145 scheme, as we will discuss in Section 4.1.

146 For non-smooth data distributions, the score function $\nabla \log q_t$ can be unbounded as $t \rightarrow 0$, which
 147 will happen especially when the data distribution is supported on a lower-dimensional submanifold
 148 of \mathbb{R}^d . This will lead to difficulty when considering the distance to the data distribution. For this
 149 reason, we consider an early stopping scheme, selecting $t_N = T - \delta$ for some small δ . Our analysis
 150 focuses on the distance between the final distribution of the sampling process and q_δ . When δ
 151 is sufficiently small, the Wasserstein-p metric between q_δ and q_0 is small. This shows that it is
 152 reasonable to approximate q_δ instead of q_0 .

153 In this paper, we propose a unified framework that can handle different time schedules. The choice
 154 of the time schedule is crucial for obtaining good results when analyzing the convergence of specific
 155 samplers. In this paper, we follow the time schedule used in Benton et al. (2024). This allows us to
 156 provide more refined control over the discretization error. We will always assume $T > 1$ and $\delta < 1$.
 157 In the first stage when $t \in [0, T-1]$, we use uniform time steps, with step size $\eta_k = t_{k+1} - t_k \leq \eta$.
 158 In the second stage for $t \in [T-1, T-\delta]$, we assume time steps satisfying $\eta_k \leq \eta(T-t_{k+1})$, which
 159 results in an exponential decay at a rate of $(1+\eta)^{-1}$. Using this schedule, we have

$$160 \quad \eta_k \leq \eta \min\{1, T-t_{k+1}\}. \quad (2.6)$$

161 It's worth noting that similar first-uniform-then-exponential schedules have also been utilized in
 Chen et al. (2023a); Li et al. (2024b;c).

162 2.2 DIFFICULTIES IN CONVERGENCE ANALYSIS OF ODE FLOW
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164 Compared with the convergence analysis of SDEs, the convergence analysis of ODEs is more chal-
 165 lenging. For SDE, the prior results either apply Girsanov’s Theorem (Chen et al., 2022) or rely on
 166 a chain-rule-based argument (Chen et al., 2023a). Using these methods, we can see that for SDEs,
 167 when the drift term is slightly perturbed, it is possible to bound the distance between the final iter-
 168 ates or, more strongly, the entire distributions over the trajectories. Therefore, the assumption of the
 169 score estimation is sufficient. However, as the following theorem illustrates, this fails in the ODE
 170 case.

170 **Theorem 2.1.** Consider the 1-dimensional OU process $(x_t)_{t \in [0, T]}$ which starts at $N(0, 1)$. It satis-
 171 fies the following SDE

$$172 \quad dx_t = -x_t dt + \sqrt{2} dW_t, \quad x_0 \sim N(0, 1).$$

174 Let the law of x_t be denoted by q_t . Its reverse process $(y_t)_{t \in [0, T]}$ (see (2.3)) can be represented with
 175 the following ODE:

$$176 \quad dy_t = \left(y_t + \nabla \log q(T-t, y_t) \right) dt, \quad y_0 \sim q_T.$$

178 For arbitrarily small $\epsilon > 0$, there exists $s_\theta(t, x)$, which is smooth, and anywhere ϵ -close to the true
 179 score function, i.e,

$$180 \quad |s_\theta(T-t, x) - \nabla \log q(T-t, x)| \leq \epsilon, \quad \forall t, \forall x,$$

182 such that the corresponding simulated reverse process $(\hat{y}_t)_{t \in [0, T]}$ (see (2.5)) satisfies

$$183 \quad TV(\hat{y}_T, y_T) \geq \frac{1}{4\pi},$$

185 which indicates that the TV-distance of the final states is larger than a constant no matter how small
 186 the score estimation error is.

187 **Remark 2.2.** For the construction of the counter-example, we consider a specific scenario where
 188 the data distribution is standard Gaussian, and the OU process maintains the distribution. In this
 189 case, adding an arbitrarily small perturbation to the drift term of the reverse ODE can result in a
 190 sampling probability density function $\hat{q}(T-t, x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} (1 + \frac{t}{2T} \sin(2n\pi x))$ with severe
 191 oscillations (n can be arbitrarily large), instead of a Gaussian distribution $q(t, x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$. For
 192 details, please refer to Section B. This demonstrates that the score estimation error assumption is
 193 insufficient, even when ignoring the discretization error. For a similar purpose, an example has been
 194 given in Li et al. (2024b) for the discrete-time sampling algorithms. Notably, our constructed s_θ
 195 is smooth while theirs is not. This indicates that adding the score function’s smoothness condition
 196 alone is insufficient to guarantee the desired properties.

198 3 CONVERGENCE OF CONTINUOUS TIME REVERSE PROCESS
 199

200 In this section, we focus on the OU forward process, where $f(t, \mathbf{x}) = -\mathbf{x}$ and $g(t) = \sqrt{2}$. Temporar-
 201 ily, we ignore the discretization error and focus on the true reverse process (2.3) and the simulated
 202 reverse process (2.5) starting at π_d . Specifically, we want to study the TV-distance between the
 203 following ODEs.

$$204 \quad d\mathbf{Y}_t = \left(\mathbf{Y}_t + \nabla \log q_{T-t}(\mathbf{Y}_t) \right) dt, \quad \mathbf{Y}_0 \sim q_T, \quad (3.1)$$

$$206 \quad d\hat{\mathbf{Y}}_t = \left(\hat{\mathbf{Y}}_t + s_\theta(T-t, \hat{\mathbf{Y}}_t) \right) dt, \quad \hat{\mathbf{Y}}_0 \sim \pi_d. \quad (3.2)$$

208 We make the following standard assumption regarding the score estimation error.

209 **Assumption 3.1** (Score estimation error). The estimated score function $s_\theta(\cdot, \cdot)$ satisfies:

$$210 \quad \mathbb{E}_{t, \mathbf{Y}_t} \|s_\theta(T-t, \mathbf{Y}_t) - \nabla \log q(T-t, \mathbf{Y}_t)\|^2 \leq \epsilon_{\text{score}}^2.$$

212 This assumption guarantees that the drift terms in (3.1) and (3.2) will be close. However, as dis-
 213 cussed in Remark 2.2, this condition alone is insufficient for ODE flows. To address this problem,
 214 we introduce the following lemma. It states that the time derivative of the TV distance between two
 215 ODE flows is determined by the distance between their drift terms and the divergence of these drift
 216 terms. Its proof is left to Appendix C.

216 **Lemma 3.2.** Suppose \mathbf{X}_t and \mathbf{Y}_t are stochastic processes in \mathbb{R}^d driven by ODEs:

$$\begin{aligned} \text{d}\mathbf{X}_t &= \mathbf{b}(t, \mathbf{X}_t) dt, \quad \mathbf{X}_0 \sim p_0, \\ \text{d}\mathbf{Y}_t &= \mathbf{b}^*(t, \mathbf{Y}_t) dt, \quad \mathbf{Y}_0 \sim q_0. \end{aligned}$$

220 Let $p(t, \mathbf{x})$ be the law of \mathbf{X}_t and $q(t, \mathbf{x})$ be the law of \mathbf{Y}_t . If the drift terms $\mathbf{b}, \mathbf{b}^* : [0, \infty) \times \mathbb{R}^d \rightarrow \mathbb{R}^d$
221 are continuously differentiable with respect to $\mathbf{x} \in \mathbb{R}^d$, then the time-derivative of the total variation
222 distance between \mathbf{X}_t and \mathbf{Y}_t satisfies the following equation:

$$\begin{aligned} \frac{\partial \text{TV}(\mathbf{X}_t, \mathbf{Y}_t)}{\partial t} &= - \int_{\Omega_t} (\nabla \cdot \mathbf{b}(t, \mathbf{x}) - \nabla \cdot \mathbf{b}^*(t, \mathbf{x})) q(t, \mathbf{x}) d\mathbf{x} \\ &\quad - \int_{\Omega_t} (\mathbf{b}(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})) \cdot \nabla \log q(t, \mathbf{x}) q(t, \mathbf{x}) d\mathbf{x}, \end{aligned}$$

228 where $\Omega_t := \{\mathbf{x} \in \mathbb{R}^d \mid p(t, \mathbf{x}) > q(t, \mathbf{x})\}$.

230 In the application of this lemma to diffusion models, we can choose the true reverse ODE and sam-
231 pling process ODE. Let p_t be the distribution of the sampling process, and q_t be the distribution
232 of the true reverse process. Several works (Chen et al., 2023b; Albergo et al., 2023), considered
233 the time derivative of KL divergence, deriving expressions that involve scores of the sampling pro-
234 cesses $\nabla \log p(t, \mathbf{x})$. However, as shown in Theorem B.1, the score of the sampling process can
235 still explode given sufficiently small score estimation error (see Assumption 3.1) and divergence
236 estimation error (see Assumption 3.3). The crucial aspect of this lemma is that it enables us to avoid
237 the occurrence of the sampling process score $\nabla \log p(t, \mathbf{x})$. We apply Gauss's theorem to transfer
238 the integral to $\partial\Omega_t$, where we can replace p_t with q_t because $p_t = q_t$ on $\partial\Omega_t$. Then, we apply
239 Gauss's theorem again to convert it back to a volume integral. This allows us to eliminate the depen-
240 dence on $\nabla \log p(t, \mathbf{x})$. Motivated by this lemma, we make the following assumption on divergence
241 estimation error.

241 **Assumption 3.3** (Divergence estimation error). For any $t \in [0, T - \delta]$, the estimated score function
242 $\mathbf{s}_\theta(t, \cdot)$ is second-order continuously differentiable. Moreover, it satisfies:

$$\mathbb{E}_{t, \mathbf{Y}_t} \left| \nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \cdot \nabla \log q(T - t, \mathbf{Y}_t) \right| \leq \epsilon_{\text{div}}.$$

245 Similar assumptions regarding the difference between the derivatives of true and estimated scores
246 have been made Li et al. (2024b;a;c). In comparison, those studies assumed the closeness of the
247 entire Jacobian matrix, while we only require assumptions on the divergence. Under Assumptions
248 3.1 and 3.3, we have the following theorem.

249 **Theorem 3.4.** Let the true and simulate reverse process be defined as (3.1) and (3.2). Under As-
250 sumptions 3.1 and 3.3, if we further assume that \mathbf{X}_0 has finite second-order momentum, then we
251 have:

$$\text{TV}(\widehat{\mathbf{Y}}_{t_N}, \mathbf{X}_\delta) \leq \text{TV}(\widehat{\mathbf{Y}}_0, \mathbf{Y}_0) + \sqrt{dT + d \log \frac{1}{\delta} \epsilon_{\text{score}}} + \epsilon_{\text{div}}.$$

255 See Appendix D for the proof of this theorem.

256 **Remark 3.5.** Huang et al. (2024a) proves a similar bound on the TV-distance for continuous-time
257 processes. Their result differs from ours in that, unlike Assumption 3.3, they do not assume the
258 divergence estimation error to be small. Instead, they require the first two derivatives of the estimated
259 score function to be bounded. Their estimation error term scales as $O(d^{3/4}T^{3/4}\delta^{-1}\epsilon_{\text{score}}^{1/2})$, which
260 is strictly worse than our $O(d^{1/2}(T^{1/2} + \log^{1/2}(\delta^{-1}))\epsilon_{\text{score}})$. Moreover, their results rely on extra
261 assumptions regarding the compact support of the data distribution. For a detailed comparison with
262 Huang et al. (2024a), regarding the settings and results, please refer to Appendix A.1.

264 4 UNIFIED ANALYSIS FOR DISCRETE-TIME REVERSE PROCESS

265 In this section, we conduct convergence analysis for the discrete-time reverse process. Specifically,
266 we consider a sequence of time steps $0 = t_0 < t_1 < \dots < t_N \leq T$. Starting from $\widehat{\mathbf{Y}}_0 \sim \pi_d$, we
267 iteratively apply a deterministic sampler $\{\mathbf{T}_k\}_{k=0}^{N-1}$ to generate subsequent iterations. For any k , the
268 sampling process can be expressed as:

$$\widehat{\mathbf{Y}}_{t_{k+1}} = \mathbf{T}_k(\widehat{\mathbf{Y}}_{t_k}),$$

270 where \mathbf{T}_k acts as a discrete-time simulation of the transition in the reverse ODE process (4.1). We
 271 will present our results in three steps. In Section 4.1, we introduce some commonly used numerical
 272 schemes for diffusion models. In Section 4.2, we provide a unified framework encompassing these
 273 numerical schemes. We then employ an interpolation method to transform the discrete-time sam-
 274 pling into an equivalent continuous-time ODE, enabling us to leverage Lemma 3.2 from the previous
 275 section. In Section 4.3, we present the convergence analysis for this general framework.

276 4.1 NUMERICAL SCHEMES

277 Recall that for general forward process (2.1), the continuous-time reverse process with the estimated
 278 score function \mathbf{s}_θ can be written as

$$280 \quad d\hat{\mathbf{Y}}_t = -\left(\mathbf{f}(T-t, \hat{\mathbf{Y}}_t) - \frac{1}{2}g(T-t)^2\mathbf{s}_\theta(T-t, \hat{\mathbf{Y}}_t)\right)dt, \quad \hat{\mathbf{Y}}_0 \sim q_T. \quad (4.1)$$

282 **Forward Euler Scheme.** The simplest method is the forward Euler scheme, which directly replaces
 283 $t \in [t_k, t_{k+1}]$ with the start point t_k in the equation above, i.e.,

$$284 \quad d\hat{\mathbf{Y}}_t = -\left(\mathbf{f}(T-t_k, \hat{\mathbf{Y}}_{t_k}) - \frac{1}{2}g(T-t_k)^2\mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k})\right)dt, \quad \text{for } t_k < t \leq t_{k+1}.$$

286 Or equivalently, we have the following discrete-time sampling algorithm, with $\eta_k = t_{k+1} - t_k$:

$$288 \quad \hat{\mathbf{Y}}_{t_{k+1}} = \hat{\mathbf{Y}}_{t_k} - \eta_k \left(\mathbf{f}(T-t_k, \hat{\mathbf{Y}}_{t_k}) - \frac{1}{2}g(T-t_k)^2\mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k}) \right). \quad (4.2)$$

290 **Exponential Integrator (EI) Scheme.** When $\mathbf{f}(t, \mathbf{y}) = L\mathbf{y}$ is a linear function, we can apply the
 291 Exponential Integrator (EI) scheme by keeping $\hat{\mathbf{Y}}_t$ in the linear part, i.e., the ODE becomes

$$293 \quad d\hat{\mathbf{Y}}_t = -\left(L\hat{\mathbf{Y}}_t - \frac{1}{2}g(T-t_k)^2\mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k})\right)dt, \quad \text{for } t_k < t \leq t_{k+1}.$$

295 In this way, we can integrate the linear part exactly. As a result, we have the following discrete-time
 296 sampling algorithm, with $\eta_k = t_{k+1} - t_k$:

$$297 \quad \hat{\mathbf{Y}}_{t_{k+1}} = e^{-L\eta_k} \hat{\mathbf{Y}}_{t_k} + \frac{e^{-L\eta_k} - 1}{2L} g(T-t_k)^2 \mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k}). \quad (4.3)$$

300 **DDIM-type Scheme.** Song et al. (2020a) introduced a deterministic sampler for the probability flow
 301 ODE by considering a non-Markovian diffusion process. As interpreted by Chen et al. (2023b), it
 302 can be viewed as a two-step process involving a restoration step that provides a rough estimate for
 303 a past step and a degradation step that simulates the forward process by progressively adding the
 304 estimated noise. Specifically, starting from $\hat{\mathbf{Y}}_{t_k}$, the restoration step provides an estimate of $\mathbf{Y}_{t_k+\gamma}$
 305 for some $\gamma > 0$, where $t_{k+1} - t_k \leq \gamma$, i.e.

$$306 \quad \mathbf{Y}_{t_k+\gamma} \approx \hat{\mathbf{Y}}_{t_k} - \gamma \left[\mathbf{f}(T-t_k, \hat{\mathbf{Y}}_{t_k}) - g(T-t_k)^2 \mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k}) \right] = \mathbf{z}.$$

308 Next, the degradation step simulates the forward process during $t \in [T-t_k-\gamma, T-t_{k+1}]$, which
 309 can be expressed as

$$310 \quad \hat{\mathbf{Y}}_{t_{k+1}} = \mathbf{z} + (t_k + \gamma - t_{k+1}) \mathbf{f}(T-t_k-\gamma, \mathbf{z}) + g(T-t_k-\gamma) \sqrt{t_k + \gamma - t_{k+1}} \boldsymbol{\epsilon},$$

312 where $\boldsymbol{\epsilon}$ represents the noise estimated from $\hat{\mathbf{Y}}_{t_k}$. By substituting the form of $\boldsymbol{\epsilon}$ and making some
 313 approximations, we can get the following sampling algorithm, with $\eta_k = t_{k+1} - t_k$:

$$315 \quad \hat{\mathbf{Y}}_{t_{k+1}} = \hat{\mathbf{Y}}_{t_k} - \eta_k \mathbf{f}(T-t_k, \hat{\mathbf{Y}}_{t_k}) + l\eta_k (1 - \sqrt{1 - 1/l}) g(T-t_k)^2 \mathbf{s}_\theta(T-t_k, \hat{\mathbf{Y}}_{t_k}), \quad (4.4)$$

316 where $l = \gamma/(t_{k+1} - t_k)$. Please refer to Chen et al. (2023b) for more details about the DDIM-type
 317 sampler. For linear diffusions, we can set $\gamma = T - t_k$. Then, with our selection of time schedule, by
 318 (2.6), we have $l \geq (T - t_{k+1})/(t_{k+1} - t_k) \geq \min\{1, T - t_{k+1}\}/(t_{k+1} - t_k) \geq 1/\eta$.

319 4.2 INTERPOLATION METHODS

321 At the core of our analysis is to apply Lemma 3.2, which analyzes the divergence between two
 322 ODEs. However, two main challenges arise. First, the discrete-time nature of the sampling algorithm
 323 precludes direct application of the lemma. Second, the sampling process $\hat{\mathbf{Y}}_{t_{k+1}} = \mathbf{T}_k(\hat{\mathbf{Y}}_{t_k})$ depends
 on the position $\hat{\mathbf{Y}}_{t_k}$ at time t_k , while the proof of Lemma 3.2 utilizes the Fokker-Planck equation,

which requires the drift term to be a function solely of time and the current position. To solve these problems, we first introduce a unified framework encompassing all the numerical schemes in Section 4.1. Next, we present an interpolation method to transform the sampling process into a continuous-time ODE, enabling the application of Lemma 3.2.

For the numerical schemes defined in (4.2), (4.3) and (4.4), we naturally extend the definition to a continuous interval $t \in [t_k, t_{k+1}]$ by replacing t_{k+1} with t . More concrete examples of F_t can be found in Section 5. This yields a continuous-time interpolation operator $F_{t_k \rightarrow t}(\cdot)$, or simply F_t when no confusion arises. Moreover, let $\hat{\mathbf{Y}}_t = F_t(\hat{\mathbf{Y}}_{t_k}) = \mathbf{F}(t, \hat{\mathbf{Y}}_{t_k})$. It is also equivalent to the following ODE:

$$d\mathbf{Y}_t = \frac{\partial F}{\partial t}(t, \hat{\mathbf{Y}}_{t_k})dt.$$

Moreover, if we further assume F_t is invertible, which holds for many examples when s_θ is Lipschitz and the time step $\eta_k = t_{k+1} - t_k$ is small enough. Then we have the following ODE:

$$d\hat{\mathbf{Y}}_t = \hat{\mathbf{b}}(t, \hat{\mathbf{Y}}_t)dt, \quad (4.5)$$

where $\hat{\mathbf{b}}(t, \mathbf{x}) = \frac{\partial F}{\partial t}(t, F_t^{-1}(\mathbf{x}))$.

4.3 MAIN RESULTS FOR GENERAL DIFFUSION PROCESSES

Using the interpolation method presented in the last section, we can now provide a general convergence analysis for the discrete-time reverse process. Consider the time step $t \in [t_k, t_{k+1}]$. Recall that for general forward process (2.1), the true reverse process is defined as

$$d\mathbf{Y}_t = \mathbf{b}(t, \mathbf{Y}_t)dt,$$

where $\mathbf{b}(t, \mathbf{x}) = -(\mathbf{f}(T-t, \mathbf{x}) - \frac{1}{2}g(T-t)^2\nabla \log q_{T-t}(\mathbf{x}))$. While the simulated reverse process is given by:

$$d\hat{\mathbf{Y}}_t = \hat{\mathbf{b}}(t, \hat{\mathbf{Y}}_t)dt,$$

where $\hat{\mathbf{b}}(t, \mathbf{x}) = \frac{\partial F}{\partial t}(t, F_t^{-1}(\mathbf{x}))$.

Definition 4.1. At each step $[t_k, t_{k+1}]$, let the interpolation operator of the sampling algorithm be $F_{t_k \rightarrow t}$. The estimation-error operator is defined by:

$$\Phi_k(t, \mathbf{x}) = \frac{\partial F}{\partial t}(t, \mathbf{x}) + \mathbf{f}(T-t, F_{t_k \rightarrow t}(\mathbf{x})) - \frac{1}{2}g(T-t)^2\nabla \log q_{T-t_k}(\mathbf{x}).$$

As we will show in the next section, it reflects the error between true score $\nabla \log q_{T-t_k}(\mathbf{x})$ and $s_\theta(T-t_k, \mathbf{x})$. Similarly, we can define the divergence-error operator, which reflects the error between $\nabla^2 \log q_{T-t_k}(\mathbf{x})$ and $\nabla s_\theta(T-t_k, \mathbf{x})$.

Definition 4.2. At each step $[t_k, t_{k+1}]$, let the interpolation operator of the sampling algorithm be $F_{t_k \rightarrow t}$. The divergence-error operator is defined by

$$\Psi_k(t, \mathbf{x}) = \nabla \left[\frac{\partial}{\partial t} F \right](t, \mathbf{x}) \mathbf{I} + \nabla_{\mathbf{x}} [\mathbf{f}(T-t, F_{t_k \rightarrow t}(\mathbf{x}))] - \frac{1}{2}g(T-t)^2 \nabla^2 \log q_{T-t_k}(\mathbf{x}).$$

Using these definitions, the next theorem shows the convergence for the general diffusion process with any numerical schemes.

Theorem 4.3. Consider the true reverse process $(\mathbf{Y}_t)_{t \in [0, T]}$ and reverse sampling process $(\hat{\mathbf{Y}}_{t_k})_{k \in [N]}$. Then,

$$\begin{aligned} \text{TV}(\mathbf{Y}_{T-\delta}, \hat{\mathbf{Y}}_{t_N}) &\leq \text{TV}(q_T, \pi_d) \\ &+ \sum_{k=0}^{N-1} \int_{t_k}^{t_{k+1}} \underbrace{\left[\sqrt{\mathbb{E}[\|\Phi_k(t, \mathbf{Y}_{t_k})\|^2]} \sqrt{\int \left\| \frac{\nabla \log q(T-t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}} \right]}_{\text{(I) Score estimation error}} \\ &+ \underbrace{\sqrt{\mathbb{E}[\text{tr}(\Psi_k(t, \mathbf{Y}_{t_k}))^2]} \sqrt{\int \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}}}_{\text{(II) Divergence estimation error}} \end{aligned}$$

$$\begin{aligned}
& + \underbrace{\frac{1}{2}g(T-t)^2 \int |(\nabla \log q_{T-t_k}(\mathbf{z}) - \nabla \log q_{T-t}(\mathbf{x})) \cdot \nabla \log q(T-t, \mathbf{x})| q(T-t, \mathbf{x}) d\mathbf{x}}_{\text{(III) Score discretization error}} \\
& + \underbrace{\frac{1}{2}g(T-t)^2 \int |\text{tr}(\nabla^2 \log q(T-t_k, \mathbf{z}) - \nabla^2 \log q(T-t, \mathbf{x}))| q(T-t, \mathbf{x}) d\mathbf{x}}_{\text{(IV) Divergence discretization error}} \\
& + \underbrace{\max_{\mathbf{x}} \left| \text{tr} \left[\nabla \left[\frac{\partial}{\partial t} F \right] (t, F_t^{-1}(\mathbf{x})) + \nabla_{\mathbf{z}} [\mathbf{f}(T-t, F_t(\mathbf{z}))] \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right|}_{\text{(V) Bias error}} dt,
\end{aligned}$$

where $\mathbf{z} = F_t^{-1}(\mathbf{x})$.

Other than the distance between the initial distribution of the reverse process and Gaussian noise, our upper bound of TV distance is divided into five terms.

For (I) and (II), these terms rely on the expectation of the defined estimation-error operator and divergence-error operator over the true reverse process. Moreover, they depend on the density ratios between $p_{Y_t}(\mathbf{x})$, representing the distribution for the true reverse process at time t , and $p_{F_t(Y_{t_k})}(\mathbf{x})$. As the interpolation operator F_t acts as a simulation of the true reverse process, we expect that the ratio will be close to 1, thus bounded. Consequently, these terms are expected to scale with the score and divergence estimation errors. For (III) and (IV), these terms depend on the distance between score functions and divergence evaluated at t and t_k , originating from the time-discretization algorithm. They will decrease as the time step gets smaller. For (V), we observe that F_t is close to the identity when $t \rightarrow t_k$. This results in $\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}$ converging to the zero matrix when $t \rightarrow t_k$.

5 APPLICATION TO SPECIFIC DETERMINISTIC SAMPLERS

In this section, we apply Theorem 4.3 to analyze specific diffusion processes. We focus on the Variance Preserving (VP) forward process with EI schemes and the Variance Exploding (VE) forward process with the DDIM sampler. While prior work has shown that VP and VE can be connected through reparametrization (Karras et al., 2022), such equivalence does not account for initialization and discretization errors. Therefore, we provide separate detailed analyses for both VP and VE processes. The analysis method can be easily extended to other forward processes and numerical schemes.

We specifically focus on data distributions with compact support, as outlined in the following assumption.

Assumption 5.1 (Bounded Support of Data). For a constant R , the data distribution q_0 satisfies:

$$q_0(\mathbf{x}) = 0, \quad \forall \|\mathbf{x}\| > R,$$

or equivalently, $P(\|\mathbf{X}_0\| > R) = 0$.

This assumption has also been made in De Bortoli (2022); Chen et al. (2022). In particular, we do not assume the smoothness of the data distribution. Therefore, it includes the setting where the data distribution is supported on a lower-dimensional submanifold of \mathbb{R}^d , which, notably, does not possess a smooth density.

Additionally, we make the following assumptions on the estimated score function.

Assumption 5.2 (Score Estimation Error).

$$\sum_k \eta_k \mathbb{E} \|\mathbf{s}_\theta(T-t_k, \mathbf{Y}_{t_k}) - \nabla \log q(T-t_k, \mathbf{Y}_{t_k})\|^2 \leq \epsilon_{\text{score}}^2.$$

Assumption 5.3 (Divergence Estimation Error).

$$\sum_k \eta_k \sqrt{\mathbb{E} \text{tr}(\nabla \mathbf{s}_\theta(T-t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q(T-t_k, \mathbf{Y}_{t_k}))^2} \leq \epsilon_{\text{div}}.$$

In addition, to deal with the discretization error for the ODE reverse process, we need the following regularity conditions for \mathbf{s}_θ .

Assumption 5.4 (\mathbf{s}_θ is Lipschitz and bounded at 0). For all t_k , we have:

$$\|\mathbf{s}_\theta(T-t_k, \mathbf{x}_1) - \mathbf{s}_\theta(T-t_k, \mathbf{x}_2)\| \leq L \cdot \|\mathbf{x}_1 - \mathbf{x}_2\|,$$

432 $\|\mathbf{s}_\theta(T - t_k, 0)\| \leq c,$
 433

434 for some constant L and c . Without loss of generality, we assume $L \geq 1$ and $c \leq L$ to simplify
 435 later derivation. Similar assumptions on \mathbf{s}_θ have been made in Huang et al. (2024a). Note that we
 436 only need the Lipschitzness and boundness at 0, while they require the boundness of high-order
 437 derivatives w.r.t t and \mathbf{x} .

438 5.1 VP+EI

439 In this section, we focus on the VP process where $\mathbf{f}(t, \mathbf{x}) = -\mathbf{x}$ and $g(t) = \sqrt{2}$. As is discussed in
 440 Section 4.1, the sampling algorithm is given by:

441 $\widehat{\mathbf{Y}}_{t_{k+1}} = e^{t_{k+1} - t_k} \widehat{\mathbf{Y}}_{t_k} + (e^{t_{k+1} - t_k} - 1) \mathbf{s}_\theta(T - t_k, \widehat{\mathbf{Y}}_{t_k}).$

443 At each step $[t_k, t_{k+1}]$, $F_t(\mathbf{x}) = e^{t-t_k} \mathbf{x} + (e^{t-t_k} - 1) \mathbf{s}_\theta(T - t_k, \mathbf{x})$. Therefore, the estimation-error
 444 operator (Definition 4.1) can be computed as

445
$$\begin{aligned}\Phi_k(t, \mathbf{x}) &= \frac{\partial F}{\partial t}(t, \mathbf{x}) + \mathbf{f}(T - t, F_t(\mathbf{x})) - \frac{1}{2}g(T - t)^2 \nabla \log q_{T-t_k}(\mathbf{x}) \\ &= s_\theta(T - t_k, \mathbf{x}) - \nabla \log q_{T-t_k}(\mathbf{x}).\end{aligned}$$

449 Similarly, the divergence-error operator (Definition 4.2) can be computed as:

450
$$\begin{aligned}\Psi_k(t, \mathbf{x}) &= \nabla \left[\frac{\partial}{\partial t} F \right](t, \mathbf{x}) \mathbf{I} + \nabla_{\mathbf{x}} [\mathbf{f}(T - t, F_{t_k \rightarrow t}(\mathbf{x}))] - \frac{1}{2}g(T - t)^2 \nabla^2 \log q_{T-t_k}(\mathbf{x}) \\ &= \nabla \mathbf{s}_\theta(T - t_k, \mathbf{x}) - \nabla^2 \log q_{T-t_k}(\mathbf{x}).\end{aligned}$$

454 Now, using Theorem 4.3, we have the following convergence analysis for VP + EI. See Appendix F
 455 for proofs of the following results.

456 **Theorem 5.5.** Consider a VP forward process with EI numerical scheme. Under Assumptions
 457 5.1, 5.2, 5.3, and 5.4, suppose the time schedule satisfies (2.6). If the step size η satisfies
 458 $\eta \leq \min\{1/(12L^2d^2), 1/(24L^2R^2d)\}$, then the following bound holds:

459
$$\text{TV}(\widehat{\mathbf{Y}}_{t_N}, \mathbf{X}_\delta) \lesssim \text{TV}(q_T, \pi_d) + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N \left[LR^4 \left(d^2 + \frac{1}{\delta^2} \right) + L^2 d + \frac{R^5}{\delta} \right].$$

462 **Lemma 5.6.** Assumes the data distribution satisfies $\text{Cov}(q_0) = \mathbf{I}_d$. Then for the VP process, we
 463 have $\text{TV}(q_T, \pi_d) \lesssim \sqrt{d} e^{-T}$. At VP forward process case, we take $\pi_d \sim N(0, \mathbf{I}_d)$.

464 **Corollary 5.7.** For all $T \geq 1$, $\delta < 1$ and $N \geq \log(1/\delta)$, there exists $\eta = \Theta((T + \log \frac{1}{\delta})/N)$
 465 and a time schedule satisfying (2.6). Under the same assumptions as Theorem 5.5, we additionally
 466 assume that the data distribution satisfies $\text{Cov}(q_0) = \mathbf{I}_d$. When $\delta = 1/d$ and $d \geq R/L + L/R^4$,
 467 if we take $T = \log(d/\epsilon^2)/2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$ for some $\epsilon \leq 1/L$, we have
 468 $\text{TV}(\widehat{\mathbf{Y}}_{t_N}, \mathbf{X}_\delta) \lesssim \epsilon$, assuming $\tilde{O}(\epsilon/\sqrt{d})$ score estimation error and $O(\epsilon)$ divergence error. Hence
 469 the diffusion model requires at most $\tilde{O}(LR^4 d^2/\epsilon)$ steps to approximate q_δ within $\epsilon \leq 1/L$ in TV
 470 distance.

471 **Remark 5.8.** We note that the ϵ -dependence in iteration complexity of our deterministic sampler is
 472 ϵ^{-1} , while stochastic ones usually exhibit a slower iteration complexity proportional to ϵ^{-2} (Chen
 473 et al., 2023a; 2022; Li et al., 2024b; Benton et al., 2024). This aligns with general observation in
 474 diffusion model practice: deterministic samplers demonstrate higher efficiency compared to stochastic
 475 samplers.

476 5.2 VE+DDIM

478 In this section, we focus on the process where $\mathbf{f}(t, \mathbf{x}) = 0$ and $g(t) = 1$, which corresponds to the
 479 VE forward process with $\sigma_t^2 = t$. As discussed in Section 4.1, the sampling algorithm is given by:

480
$$\widehat{\mathbf{Y}}_{t_{k+1}} = \widehat{\mathbf{Y}}_{t_k} + (t_{k+1} - t_k)l(1 - \sqrt{1 - 1/l}) \mathbf{s}_\theta(T - t_k, \widehat{\mathbf{Y}}_{t_k}), \quad l = \frac{T - t_k}{t_{k+1} - t_k}.$$

483 Let $c_l = l(1 - \sqrt{1 - 1/l}) \leq 1$. At each step $[t_k, t_{k+1}]$, $F_t(\mathbf{x}) = \mathbf{x} + c_l(t - t_k) \mathbf{s}_\theta(T - t_k, \mathbf{x})$.
 484 Therefore, the estimation-error operator (Definition 4.1) can be computed as

485
$$\Phi_k(t, \mathbf{x}) = \frac{\partial F}{\partial t}(t, \mathbf{x}) + \mathbf{f}(T - t, F_t(\mathbf{x})) - \frac{1}{2}g(T - t)^2 \nabla \log q_{T-t_k}(\mathbf{x})$$

$$= c_l s_\theta(T - t_k, \mathbf{x}) - \frac{1}{2} \nabla \log q_{T-t_k}(\mathbf{x}).$$

Similarly, the divergence-error operator (Definition 4.2) can be computed as:

$$\begin{aligned} \Psi_k(t, \mathbf{x}) &= \nabla \left[\frac{\partial}{\partial t} F \right] (t, \mathbf{x}) \mathbf{I} + \nabla_{\mathbf{x}} [\mathbf{f}(T - t, F_{t_k \rightarrow t}(\mathbf{x}))] - \frac{1}{2} g(T - t)^2 \nabla^2 \log q_{T-t_k}(\mathbf{x}) \\ &= c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{x}) - \frac{1}{2} \nabla^2 \log q_{T-t_k}(\mathbf{x}). \end{aligned}$$

Now, using Theorem 4.3, we have the following convergence analysis for VE + DDIM. See Appendix G for proofs of the following results.

Theorem 5.9. Consider a VE forward process with DDIM numerical scheme. Under Assumptions 5.1, 5.2, 5.3, and 5.4, suppose the time schedule satisfies (2.6). If the step size η satisfies $\eta \leq \min\{1/(12L^2Td^2), 1/(24L^2R^2d)\}$, then the following bound holds:

$$\text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \text{TV}(\pi_d, q_T) + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N \left[LR^4 d^2 + L^2 d + \frac{LR^4}{\delta^2} \right].$$

Lemma 5.10. Assume the data distribution satisfies $\text{Cov}(q_0) = \mathbf{I}_d$. Then for the VE forward process, we have $\text{TV}(q_T, \pi_d) \lesssim \sqrt{d}/\sqrt{T}$. At VE forward process case, we take $\pi_d \sim N(0, T\mathbf{I}_d)$.

Corollary 5.11. For all $T \geq 1$, $\delta < 1$ and $N \geq \log(1/\delta)$, there exists $\eta = \Theta((T + \log \frac{1}{\delta})/N)$ and a time schedule satisfying (2.6). Under the same assumptions as Theorem 5.5, we additionally assume that the data distribution satisfies $\text{Cov}(q_0) = \mathbf{I}_d$. When $\delta = 1/d$ and $d \geq L/R^4$, if we take $T = d/\epsilon^2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$ for some $\epsilon \leq 1/L$, we have $\text{TV}(\widehat{\mathbf{Y}}_{t_N}, \mathbf{X}_\delta) \lesssim \epsilon$, assuming sufficiently small score estimation error and divergence error. Hence the diffusion model requires at most $\widetilde{O}(LR^4 d^4/\epsilon^5)$ steps to approximate q_δ within $\epsilon \leq 1/L$ in TV distance.

Remark 5.12. Compared to Corollary 5.7, VE SDE has larger iteration complexity, primarily due to the slow $1/\sqrt{T}$ decay in the distance between q_T and π_d , compared with the exponential decay (Lemma 5.6) for VP SDE.

6 CONCLUSION

In this work, we introduce a unified convergence analysis framework for deterministic samplers. We start by presenting a counter-example to illustrate the main challenge in analyzing deterministic samplers compared to stochastic ones. Additionally, we provide a technical lemma that allows us to bound the distance between distributions using score estimation error and divergence error. With this approach, we directly established convergence guarantees for the continuous-time reverse ODE. Moreover, we extend our analysis to the convergence of discrete-time deterministic samplers with a unified framework. Finally, we demonstrate its effectiveness by applying it to two widely adopted sampling methods.

Limitation and Future Work. First, for the VP process with EI schemes, our current results have a quadratic dependence on d , which leaves room for improvement compared to the d -linear state-of-the-art bounds in ODE analysis (Li et al., 2024c). This discrepancy stems from two factors: our different assumption that requires control of the divergence error rather than the full Jacobian error, and our directly estimated Lipschitz constants in the discretization error analysis. We believe that through more delicate methods for analyzing the discretization error and potentially stronger conditions, we could achieve linear dimension dependence within our unified framework. Second, since adding the divergence assumption can guarantee the convergence of the ODE sampler, an interesting future direction may be to design training methods that can obtain both small score error and small divergence error. One potential approach would be to incorporate regularization terms corresponding to divergence error in the loss function, potentially leading to more effective diffusion model training algorithms. Third, for the VE process, we obtain only polynomial bounds due to the slow $1/\sqrt{T}$ decay in the distance between q_T and π_d . It's important to note that this is not a limitation of the DDIM sampler, as applying DDIM to the VP processes yields results comparable to those obtained with the EI scheme. We leave an improved analysis for the VE forward process as future work. Finally, the discrete-time analysis currently relies on a bounded support assumption for the data distribution, which may be relaxed to less restrictive conditions, such as light-tailed distributions.

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702 **A RELATED WORK**

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 704 **Convergence Analysis of Stochastic Samplers.** Early theoretical studies for diffusion models were
 705 either non-quantitative (De Bortoli et al., 2021; Liu et al., 2022; Pidstrigach, 2022), or exhibited ex-
 706 ponential dependence on the dimension or other problem parameters (Block et al., 2020; De Bortoli,
 707 2022). Later, Lee et al. (2022) proved the first result with polynomial complexity, with the assump-
 708 tions L_2 -score estimate and log-Sobolev inequality (LSI). However, the LSI condition on the data
 709 distribution is restrictive, prompting further studies to relax this assumption. Chen et al. (2022)
 710 utilized Girsanov’s theorem, and proved polynomial convergence bounds, assuming either the Lips-
 711 chitzness property of the forward process score function or bounded support of the data distribution.
 712 Lee et al. (2023) relaxed the smoothness condition to apply only to the data distribution rather
 713 than the whole trajectory, and replaced the bounded support assumption with sufficient tail decay.
 714 Chen et al. (2023a) proved a result with both the advantages of these two works, achieving a better
 715 convergence rate while assuming only the smoothness of the data distribution. Furthermore, they
 716 provided results for the non-smooth setting using appropriate early stopping and decreasing step
 717 size. Benton et al. (2024) improved the dependency of the dimension d to linear via the stochas-
 718 tic localization method. Compared with the analysis of reverse SDE, Li et al. (2024b) utilized an
 719 elementary approach to analyze a DDPM-type stochastic sampler. It also proposed an accelerated
 720 stochastic sampler with better iteration complexity. More recently, Huang et al. (2024b) decom-
 721 posed the entire sampling process into several reverse transition kernel subproblems and proposed
 722 novel fast sampling algorithms.

723 **Convergence Analysis of Deterministic Samplers.** The first non-asymptotic bounds for deter-
 724 ministic samplers were derived by Chen et al. (2023b), which assumes access to the ground truth score
 725 function. Subsequently, Chen et al. (2024) proved the first polynomial-time convergence guarantees
 726 of probability flow ODE with estimation error, by incorporating an additional corrector Langevin
 727 dynamics. While this improved upon prior results, it introduced randomness, making the sampling
 728 processes non-deterministic. Li et al. (2024a;b) applied an elementary analysis framework to study
 729 the convergence of a specific deterministic sampler with a given learning schedule, requiring an ad-
 730 ditional assumption on the Jacobian estimation error. Later, Li et al. (2024c) applied the same frame-
 731 work and achieved improved iteration complexity, which is linear in the dimension. Recently, Huang
 732 et al. (2024a) examined the Ornstein-Uhlenbeck (OU) forward process, removing the assumption of
 733 Jacobian estimation error. However, they introduced higher-order boundedness assumptions on the
 734 derivatives of the estimated score function concerning time and space. They proved an upper bound
 735 for the total variation between the target and generated data at the continuous-time level while also
 736 analyzing the convergence rate for the Runge-Kutta integrator. Further expanding the scope, Gao
 737 & Zhu (2024) studied the Wasserstein distance using general forward SDE with log-concavity data
 738 assumptions. Additionally, research has started to explore other deterministic generation methods,
 739 such as flow matching (Benton et al., 2023).

740 A.1 COMPARISON WITH HUANG ET AL. (2024A)

741 Huang et al. (2024a) is the most related work to ours. It studied the convergence properties of
 742 diffusion models with deterministic samplers based on probability flow ODEs.

743 Theorem A.1 in Huang et al. (2024a) is similar to our Lemma 3.2, where they applied the charac-
 744 teristic line method for ODEs, while we used the Fokker-Planck equation and the Gauss’s theorem
 745 twice. However, we use different methods when dealing with the divergence term

$$\nabla \cdot [((\mathbf{b}(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})) q(T-t, \mathbf{x})].$$

746 We decompose it into score estimation error and divergence estimation error (see Lemma 3.2). In
 747 contrast, they applied the Gagliardo-Nirenberg inequality to bound the integral of first-order deriva-
 748 tives using integrals of second-order and zero-order derivatives of \mathbf{s}_θ . This led to different assump-
 749 tions regarding the convergence analysis of the continuous-time reverse process. In addition to
 750 the standard estimation error assumption, we require a small divergence error, while Huang et al.
 751 (2024a) assumed the boundedness of up to the second order derivatives of \mathbf{s}_θ . Furthermore, they re-
 752 quired a bounded support assumption, which is not necessary for our continuous-time analysis. As
 753 a result, Huang et al. (2024a) proved an upper bound on the TV distance: $O(T^{3/4}d^{3/4}(1/\delta)^{1/2}\epsilon_{\text{score}})$.
 754 In comparison, our result is $O(\sqrt{dT + d \log(1/\delta)} \cdot \epsilon_{\text{score}})$, notably only log-dependent on δ .

755 For the discrete-time analysis, Huang et al. (2024a) studied the p -th Runge-Kutta discretization
 756 method. When $p = 1$, it becomes the forward Euler method. In comparison, we study the EI
 757 scheme. Our analysis can be easily applied to the forward Euler method, yielding similar results.

In the discrete-time case, we assumed bounded support of the data distribution and the Lipschitz condition of \mathbf{s}_θ , which was already required in Huang et al. (2024a) in the continuous-time analysis. Additionally, they assumed that the mixed second-order derivatives of \mathbf{s}_θ with respect to t and x are bounded by a constant \bar{L} . For the forward Euler method, Huang et al. (2024a) proved iteration complexity of $\tilde{O}(\bar{L}^2 R^2 d^2 \epsilon^{-1})$, while our result for EI scheme is $\tilde{O}(LR^4 d^2 \epsilon^{-1})$ for $\delta = 1/d$. We remark that the constant \bar{L} in their result may be very large. For example, the derivatives of the true score function can depend on higher-order terms of $1/\delta$ (see Lemma H.2). Moreover, we also consider the VE forward process with the DDIM numerical schemes.

A.2 COMPARISON WITH OTHER WORKS ON DETERMINISTIC SAMPLERS

In this section, we consider different works on convergence analysis of diffusion models with deterministic samplers and provide a detailed comparison in Table 1. The comparison focuses on three key aspects: major assumptions required by these works, the metrics used, and the iteration complexity in terms of dimension d and accuracy ϵ .

Major Assumptions	Metric	Complexity	Reference
Score Estimation Error ¹		$\tilde{O}\left(\frac{d^2}{\epsilon} + \frac{d^3}{\sqrt{\epsilon}}\right)$	Li et al. (2024b)
Bounded Support	$\text{TV}(p_1, q_1)$		Theorem 1
$\frac{1}{N} \sum_{t=1}^N \mathbb{E}_{\mathbf{X} \sim q_t} \ \nabla \mathbf{s}_\theta(t, \mathbf{X}) - \nabla^2 \log q(t, \mathbf{X})\ \leq \epsilon_{\text{Jacobi}}$		$\tilde{O}\left(\frac{d}{\epsilon}\right)$	Li et al. (2024c) Theorem 1
Access to Exact Score $\nabla \log q$ and $\nabla^2 \log q$ Lipschitz $\ \nabla \log \frac{q_t}{q_s}(\mathbf{x})\ \leq \beta t-s ^c (1 + \ \mathbf{x}\ + \ \nabla q_t(\mathbf{x})\)$	$\text{KL}(p_0, q_0)$	Polynomial	Chen et al. (2023b) Theorem 4.1
Score Estimation Error $\sup_{\mathbf{x} \in \mathbb{R}^d} \max_{1 \leq k+ \alpha \leq p+1} \max_{1 \leq j \leq d} \partial_t^k \partial_\mathbf{x}^\alpha \mathbf{s}_\theta^{(j)}(t, \mathbf{x}) \leq L$	$\text{TV}(p_\delta, q_\delta)$	$O\left(\frac{d^{(p+1)/p}}{\epsilon^{1/p}}\right)$	Huang et al. (2024a) Theorem 3.10
Bounded Support: Assumption 5.1 Score Estimation Error: Assumption 5.2 Divergence Error: Assumption 5.3	$\text{TV}(p_\delta, q_\delta)$	$\tilde{O}\left(\frac{d^2}{\epsilon}\right)$	This Work: Corollary 5.7

Table 1: Comparison of convergence analysis for diffusion models with deterministic samplers

B PROOF OF THEOREM 2.1

Proof of Theorem 2.1. Since x_t starts at $N(0, 1)$, we can easily show that $x_t \sim N(0, 1)$ for all t . Recall that we use $q(t, x)$ to denote the law of x_t , then we have $q(t, x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$ and $\nabla \log q(t, x) = -x$. Define $\hat{q}(T-t, x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} (1 + \frac{t}{2T} \sin(2n\pi x))$, and construct our estimated score as:

$$\mathbf{s}_\theta(T-t, x) = \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \sin(2n\pi y) \left(\frac{1}{2T}\right) dy}{\hat{q}(T-t, x)} - x,$$

Without loss of generality, we can assume $x \geq 0$. Then we have:

$$|\mathbf{s}_\theta(T-t, x) - \nabla \log q(T-t, x)| = \left| \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \sin(2n\pi y) \left(\frac{1}{2T}\right) dy}{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} (1 + \frac{T-t}{2T} \sin(2n\pi x))} - x \right|$$

¹This assumption is slightly different from ours. While it assigns uniform weights to estimation errors across all time steps, our approach applies smaller weights to time steps with smaller step sizes.

$$\begin{aligned}
&\leq \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{2T} \sup_{N \geq x} \left| \int_x^N \sin(2n\pi y) dy \right|}{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{2}} \\
&\leq \frac{1}{T} \frac{1}{n\pi}.
\end{aligned}$$

Taking $n \geq 1/(T\pi\epsilon)$, we have $|\mathbf{s}_\theta(T-t, x) - \nabla \log q(T-t, x)| \leq \epsilon$. Moreover, since we have

$$\frac{\partial \hat{q}(T-t, x)}{\partial t} = -\frac{\partial}{\partial x} \left((\mathbf{s}_\theta(T-t, x) + x) \cdot \hat{q}(T-t, x) \right), \quad \hat{q}(T-0, x) \sim N(0, 1),$$

which satisfies the Fokker-Planck equation. Therefore, we know that $\hat{q}(t, x)$ is the law of \hat{y}_{T-t} . Hence the law of \hat{y}_T is

$$\hat{q}_0(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(1 + \frac{1}{2} \sin(2n\pi x) \right).$$

The TV distance between y_T and \hat{y}_T is:

$$\begin{aligned}
\text{TV}(y_T, \hat{y}_T) &= \frac{1}{2} \int_{\mathbb{R}} \left| \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(1 + \frac{1}{2} \sin(2n\pi x) \right) - \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \right| dx \\
&= \frac{1}{4} \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} |\sin(2n\pi x)| dx. \tag{B.1}
\end{aligned}$$

To calculate the last integral, we consider the Fourier expansion

$$|\sin(2n\pi x)| = \frac{2}{\pi} - \frac{4}{\pi} \sum_{k \geq 1} \frac{\cos 4kn\pi x}{4k^2 - 1}.$$

For any $k \geq 1$, we have:

$$\begin{aligned}
\left| \int_{\mathbb{R}} e^{-\frac{x^2}{2}} \cos(4kn\pi x) dx \right| &= 2 \left| \int_0^\infty e^{-\frac{x^2}{2}} \cos(4kn\pi x) dx \right| \\
&\leq 2 \sup_{N \geq 0} \left| \int_0^N \cos(4kn\pi x) dx \right| \\
&\leq \frac{1}{2kn\pi}.
\end{aligned}$$

Therefore, we can plug these integrals into B.1 and obtain

$$\begin{aligned}
\text{TV}(y_T, \hat{y}_T) &\geq \frac{1}{4} \left[\frac{2}{\pi} - \frac{4}{\pi} \sum_{k \geq 1} \frac{1}{4k^2 - 1} \left| \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cos(4kn\pi x) dx \right| \right] \\
&\geq \frac{1}{4} \left[\frac{2}{\pi} - \frac{4}{\pi} \sum_{k \geq 1} \frac{1}{4k^2 - 1} \frac{1}{\sqrt{2\pi} 2kn\pi} \right] \\
&\geq \frac{1}{4} \left[\frac{2}{\pi} - \frac{2}{\pi^2 \sqrt{2\pi n}} \sum_{k \geq 1} \frac{1}{3k^3} \right].
\end{aligned}$$

By taking n large enough, we can easily show that $\text{TV}(y_T, \hat{y}_T) \geq \frac{1}{4\pi}$. This completes the proof of Theorem 2.1. \square

Theorem B.1. Consider the same 1-dimensional OU process $(x_t)_{t \in [0, T]}$ and its reverse process $(y_t)_{t \in [0, T]}$ defined in Theorem 2.1. Recall that we denote the law of x_t by q_t . Then for arbitrarily small $\epsilon > 0$ and arbitrarily big $N > 0$, there exists $s_\theta(t, x)$, which is smooth, anywhere ϵ -close to the true score function, and its divergence is anywhere ϵ -close to the divergence of the true score function, i.e.,

$$\begin{aligned}
|\mathbf{s}_\theta(T-t, x) - \nabla \log q(T-t, x)| &\leq \epsilon, \quad \forall t, \forall x, \\
|\nabla \cdot \mathbf{s}_\theta(T-t, x) - \nabla \cdot \nabla \log q(T-t, x)| &\leq \epsilon, \quad \forall t, \forall x,
\end{aligned}$$

such that the law p_t of the corresponding simulated reverse process $(\hat{y}_t)_{t \in [0, T]}$ (see (2.5)) satisfies

$$\sup |\nabla \log p_T(x)| \geq N,$$

which indicates that the score of the sampling process can be unbounded no matter how small the score estimation error and divergence error is.

864 *Proof of Theorem B.1.* Substitute the oscillate term $\sin(2n\pi x)$ in the proof of Theorem 2.1 by
 865 $\frac{1}{n(x^2+1)} \sin(2n^2\pi x)$. In this case, we still have $q(t, x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$ and $\nabla \log q(t, x) = -x$. Now
 866 $\hat{q}(T-t, x)$ becomes $\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(1 + \frac{t}{2T} \frac{1}{n(x^2+1)} \sin(2n^2\pi x) \right)$, and now our estimated score becomes:
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$$869 \quad s_\theta(T-t, x) = \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \left(\frac{1}{2T} \right) dy}{\hat{q}(T-t, x)} - x,$$

870 Without loss of generality, we can assume $x \geq 0$. Then we have:
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 872

$$\begin{aligned} 873 \quad |s_\theta(T-t, x) - \nabla \log q(T-t, x)| &= \left| \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \left(\frac{1}{2T} \right) dy}{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(1 + \frac{T-t}{2T} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \right)} \right. \\ 874 \quad &\leq \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{2T} \frac{1}{(y^2+1)} \sup_{N \geq x} \left| \int_x^N \frac{1}{n} \sin(2n^2\pi y) dy \right|}{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{2}} \\ 875 \quad &\leq \frac{1}{T} \frac{1}{n^3 \pi}. \end{aligned} \tag{B.2}$$

876 Further more, we consider the divergence estimation error. Since this example is one-dimensional,
 877 taking divergence is the same as taking a derivative with respect to x . Hence we have:
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$$\begin{aligned} 880 \quad |\nabla \cdot s_\theta(T-t, x) - \nabla \cdot \nabla \log q(T-t, x)| &= \left| \frac{\partial}{\partial x} \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \left(\frac{1}{2T} \right) dy}{\hat{q}(T-t, x)} \right| \\ 881 \quad &\leq \left| \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{n(x^2+1)} \sin(2n^2\pi x) \left(\frac{1}{2T} \right) \cdot \hat{q}(T-t, x)}{\hat{q}(T-t, x)^2} \right| \\ 882 \quad &\quad + \left| \frac{\int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \left(\frac{1}{2T} \right) dy \cdot \frac{\partial}{\partial x} \hat{q}(T-t, x)}{\hat{q}(T-t, x)^2} \right|. \end{aligned} \tag{B.3}$$

883 Firstly, we have
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$$886 \quad \frac{1}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \leq |\hat{q}(T-t, x)| \leq \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}.$$

887 Moreover, we know that
 888
 889

$$\begin{aligned} 890 \quad \left| \frac{\partial}{\partial x} \hat{q}(T-t, x) \right| &= \left| \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(-x \left(1 + \frac{t}{2T} \frac{1}{n(x^2+1)} \sin(2n^2\pi x) \right) \right. \right. \\ 891 \quad &\quad \left. \left. + \frac{t}{2T} \frac{1}{n} \frac{2n^2\pi \cos(2n^2\pi x)(x^2+1) - 2x \sin(2n^2\pi x)}{(x^2+1)^2} \right) \right| \\ 892 \quad &\leq \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \left(|x| + \frac{1}{4Tn} + n\pi + \frac{1}{2Tn} \right) \\ 893 \quad &\leq \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} (|x| + 2n\pi). \end{aligned} \tag{B.4}$$

894 Since $\frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)}$ is strictly monotonically decreasing when $y \geq x \geq 0$, we have
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$$\begin{aligned} 897 \quad &\left| \int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \frac{1}{n(y^2+1)} \sin(2n^2\pi y) \left(\frac{1}{2T} \right) dy \right| \\ 898 \quad &\leq \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{n(x^2+1)} \frac{1}{2T} \sup_{N \geq x} \left| \int_x^N \sin(2n^2\pi y) dy \right| \\ 899 \quad &\leq \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{x^2+1} \frac{1}{n^3 \pi} \frac{1}{2T}. \end{aligned}$$

918 Substituting back to B.3, we have
 919

$$\begin{aligned} 920 \quad |\nabla \cdot \mathbf{s}_\theta(T-t, x) - \nabla \cdot \nabla \log q(T-t, x)| &\leq \frac{1}{2Tn} + 4 \frac{1}{x^2+1} \frac{1}{n^3\pi} \frac{1}{2T} (|x| + 2n\pi) \\ 921 \quad &\leq \frac{1}{2Tn} + \frac{1}{Tn^3\pi} + \frac{4}{Tn^2}. \end{aligned} \quad (\text{B.5})$$

924 For any $\epsilon > 0$, using (B.2) and (B.5), for n large enough, the score estimation error and the divergence estimation error are less than ϵ for every t and x . Same as in the proof of Theorem 2.1, using Fokker-Planck equation, we know that \hat{q} is the law of the sampling process. However, by (B.4), the score of the sampling process is clearly unbounded (there is a coefficient proportional to n , and our n is arbitrary large). \square

930 C PROOF OF LEMMA 3.2

932 *Proof of Lemma 3.2.* We begin by computing total variance distance between \mathbf{X}_t and \mathbf{Y}_t for any
 933 $t \in [0, T]$:

$$935 \quad \text{TV}(\mathbf{X}_t, \mathbf{Y}_t) = \int_{\Omega_t} p(t, \mathbf{x}) - q(t, \mathbf{x}) d\mathbf{x}, \quad (\text{C.1})$$

937 where $\Omega_t := \{\mathbf{x} \in \mathbb{R}^d \mid p(t, \mathbf{x}) > q(t, \mathbf{x})\}$. Now we proceed to compute the time-derivative
 938 of (C.1). By Theorem K.1, we have
 939

$$\begin{aligned} 940 \quad \frac{\partial \text{TV}(\mathbf{X}_t, \mathbf{Y}_t)}{\partial t} &= \underbrace{\int_{\Omega_t} \frac{\partial}{\partial t} (p(t, \mathbf{x}) - q(t, \mathbf{x})) d\mathbf{x}}_{I_1} + \underbrace{\int_{\partial\Omega_t} (p(t, \mathbf{x}) - q(t, \mathbf{x})) \mathbf{v}(t, \mathbf{x}) \cdot \mathbf{n}(t, \mathbf{x}) dS}_{I_2}. \\ 941 \quad & \end{aligned} \quad (\text{C.2})$$

945 For term I_1 , by Fokker-Planck equation, we have
 946

$$947 \quad \frac{\partial}{\partial t} (p(t, \mathbf{x}) - q(t, \mathbf{x})) = -\nabla \cdot (\mathbf{b}(t, \mathbf{x})p(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})q(t, \mathbf{x})).$$

949 Therefore, we have
 950

$$951 \quad I_1 = \int_{\Omega_t} -\nabla \cdot (\mathbf{b}(t, \mathbf{x})p(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})q(t, \mathbf{x})) d\mathbf{x}.$$

954 For term I_2 , by the definition of $\Omega_t := \{\mathbf{x} \in \mathbb{R}^d \mid p(t, \mathbf{x}) > q(t, \mathbf{x})\}$, we have $p(t, \mathbf{x}) = q(t, \mathbf{x})$ on
 955 $\partial\Omega_t$. Hence $I_2 = 0$.

956 Therefore, by substituting I_1 and I_2 into (C.2), we obtain
 957

$$\begin{aligned} 958 \quad \frac{\partial \text{TV}(\mathbf{X}_t, \mathbf{Y}_t)}{\partial t} &= \int_{\Omega_t} -\nabla \cdot (\mathbf{b}(t, \mathbf{x})p(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})q(t, \mathbf{x})) d\mathbf{x} \\ 959 \quad &= \int_{\partial\Omega_t} -(\mathbf{b}(t, \mathbf{x})p(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})q(t, \mathbf{x})) \cdot \mathbf{n}(t, \mathbf{x}) dS \\ 960 \quad &= \int_{\partial\Omega_t} -(\mathbf{b}(t, \mathbf{x})q(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})q(t, \mathbf{x})) \cdot \mathbf{n}(t, \mathbf{x}) dS \\ 961 \quad &= \int_{\Omega_t} -\nabla \cdot ((\mathbf{b}(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x}))q(t, \mathbf{x})) d\mathbf{x} \\ 962 \quad &= - \int_{\Omega_t} (\nabla \cdot \mathbf{b}(t, \mathbf{x}) - \nabla \cdot \mathbf{b}^*(t, \mathbf{x}))q(t, \mathbf{x}) d\mathbf{x} - \int_{\Omega_t} (\mathbf{b}(t, \mathbf{x}) - \mathbf{b}^*(t, \mathbf{x})) \cdot \nabla q(t, \mathbf{x}) d\mathbf{x}, \end{aligned}$$

970 where in the second equation we use Gauss's theorem to compute the integration of divergence,
 971 the third equation uses the fact that $p(t, \mathbf{x}) = q(t, \mathbf{x})$ on $\partial\Omega_t$, the forth equation holds by Gauss's
 972 theorem again. This completes the proof. \square

972 **D PROOF OF THEOREM 3.4**
 973

974 In this section, we provide a proof of Theorem 3.4. The core of the theorem’s proof is Lemma 3.2,
 975 which allows us to represent the time-derivative of TV distribution between distribution by score
 976 estimation error and divergence estimation error.

977 *Proof of Theorem 3.4.* Recall \mathbf{Y}_t and $\widehat{\mathbf{Y}}_t$ are ODE flows determined by (3.1) and (3.2). Define
 978 $\mathbf{b}(t, \mathbf{x}) = \mathbf{x} + \nabla \log q(T - t, \mathbf{x})$ and $\widehat{\mathbf{b}}(t, \mathbf{x}) = \mathbf{x} + \mathbf{s}_\theta(T - t, \mathbf{x})$, then we have:
 979

$$\begin{aligned} d\mathbf{Y}_t &= \mathbf{b}(t, \mathbf{Y}_t)dt, \quad \mathbf{Y}_0 \sim q_T. \\ d\widehat{\mathbf{Y}}_t &= \widehat{\mathbf{b}}(t, \widehat{\mathbf{Y}}_t)dt, \quad \widehat{\mathbf{Y}}_0 \sim \pi_d. \end{aligned}$$

982 By Lemma 3.2, recall that the law of \mathbf{Y}_t is denoted by $q(T - t, \mathbf{x})$, we have:
 983

$$\begin{aligned} \frac{\partial \text{TV}(\widehat{\mathbf{Y}}_t, \mathbf{Y}_t)}{\partial t} &= - \int_{\Omega_t} (\nabla \cdot \widehat{\mathbf{b}}(t, \mathbf{x}) - \nabla \cdot \mathbf{b}(t, \mathbf{x})) q(T - t, \mathbf{x}) d\mathbf{x} \\ &\quad - \int_{\Omega_t} (\widehat{\mathbf{b}}(t, \mathbf{x}) - \mathbf{b}(t, \mathbf{x})) \cdot \nabla \log q(T - t, \mathbf{x}) q(T - t, \mathbf{x}) d\mathbf{x} \\ &= \underbrace{\int_{\Omega_t} (\nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{x}) - \nabla \cdot \nabla \log q(T - t, \mathbf{x})) q(T - t, \mathbf{x}) d\mathbf{x}}_{I_1} \\ &\quad - \underbrace{\int_{\Omega_t} (\mathbf{s}_\theta(T - t, \mathbf{x}) - \nabla \log q(T - t, \mathbf{x})) \cdot \nabla \log q(T - t, \mathbf{x}) q(T - t, \mathbf{x}) d\mathbf{x}}_{I_2}. \end{aligned} \tag{D.1}$$

987 For I_1 , we know that:
 988

$$\begin{aligned} |I_1| &\leq \int_{\Omega_t} |\nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{x}) - \nabla \cdot \nabla \log q(T - t, \mathbf{x})| q(T - t, \mathbf{x}) d\mathbf{x} \\ &\leq \mathbb{E} |\nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \cdot \nabla \log q(T - t, \mathbf{Y}_t)|. \end{aligned} \tag{D.2}$$

992 For I_2 , using the Cauchy-Schwartz inequality, we know that:
 993

$$\begin{aligned} |I_2| &\leq \sqrt{\int_{\Omega_t} \|\mathbf{s}_\theta(T - t, \mathbf{x}) - \nabla \log q(T - t, \mathbf{x})\|^2 q(T - t, \mathbf{x}) d\mathbf{x}} \\ &\quad \cdot \sqrt{\int_{\Omega_t} \|\nabla \log q(T - t, \mathbf{x})\|^2 q(T - t, \mathbf{x}) d\mathbf{x}} \\ &\leq \sqrt{\mathbb{E} \|\mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \log q(T - t, \mathbf{Y}_t)\|^2} \cdot \sqrt{\mathbb{E} \|\nabla \log q(T - t, \mathbf{Y}_t)\|^2}. \end{aligned} \tag{D.3}$$

1011 Substituting (D.2) and (D.3) back to (D.1), and taking the integral with respect to t on $[0, T - \delta]$, we
 1012 have:
 1013

$$\begin{aligned} \text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{X}_\delta) &\leq \text{TV}(\widehat{\mathbf{Y}}_0, \mathbf{Y}_0) + \int_0^{T-\delta} (|I_1| + |I_2|) dt \\ &\leq \text{TV}(\widehat{\mathbf{Y}}_0, \mathbf{Y}_0) + \int_0^{T-\delta} \mathbb{E} |\nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \cdot \nabla \log q(T - t, \mathbf{Y}_t)| dt \\ &\quad + \int_0^{T-\delta} \sqrt{\mathbb{E} \|\mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \log q(T - t, \mathbf{Y}_t)\|^2} \cdot \sqrt{\mathbb{E} \|\nabla \log q(T - t, \mathbf{Y}_t)\|^2} dt \\ &\leq \text{TV}(\widehat{\mathbf{Y}}_0, \mathbf{Y}_0) + \int_0^{T-\delta} \mathbb{E} |\nabla \cdot \mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \cdot \nabla \log q(T - t, \mathbf{Y}_t)| dt \\ &\quad + \sqrt{\int_0^{T-\delta} \mathbb{E} \|\mathbf{s}_\theta(T - t, \mathbf{Y}_t) - \nabla \log q(T - t, \mathbf{Y}_t)\|^2 dt} \cdot \sqrt{\int_0^{T-\delta} \mathbb{E} \|\nabla \log q(T - t, \mathbf{Y}_t)\|^2 dt}. \end{aligned} \tag{D.4}$$

1026 Since \mathbf{X}_0 has finite second-order momentum, using Lemma K.2, we have that:
1027

$$\mathbb{E}_Q \|\nabla \log q(T-t, \mathbf{Y}_t)\|^2 \leq d(1 - e^{-2(T-t)})^{-1}.$$

1029 Thus we know that:
1030

$$\begin{aligned} \int_0^{T-\delta} \mathbb{E} \|\nabla \log q(T-t, \mathbf{Y}_t)\|^2 dt &\leq \int_0^{T-\delta} \frac{d}{1 - e^{-2(T-t)}} dt \\ &= d \int_{\delta}^T \frac{1}{1 - e^{-2t}} dt \\ &= d \left(\frac{1}{2} \log(e^{2T} - 1) - \frac{1}{2} \log(e^{2\delta} - 1) \right) \\ &= O\left(dT + d \log \frac{1}{\delta}\right). \end{aligned} \quad (\text{D.5})$$

1040 Substituting (D.5) into (D.4) and using Assumptions 3.1 and 3.3, we can conclude with
1041

$$\text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{X}_\delta) \lesssim \text{TV}(\widehat{\mathbf{Y}}_0, \mathbf{Y}_0) + \sqrt{dT + d \log \frac{1}{\delta}} \epsilon_{\text{score}} + \epsilon_{\text{div}}.$$

1044 \square
1045

1046 E PROOF OF THEOREM 4.3

1047 *Proof of Theorem 4.3.* In the time step $t \in [t_k, t_{k+1}]$, we can compute the time derivative of the
1048 total variation distance between $\widehat{\mathbf{Y}}_t$ and \mathbf{Y}_t using Lemma 3.2:

$$\begin{aligned} \frac{\partial \text{TV}(\widehat{\mathbf{Y}}_t, \mathbf{Y}_t)}{\partial t} &= - \underbrace{\int_{\Omega_t} (\nabla \cdot \widehat{\mathbf{b}}(t, \mathbf{x}) - \nabla \cdot \mathbf{b}(t, \mathbf{x})) q(T-t, \mathbf{x}) d\mathbf{x}}_{I_1} \\ &\quad - \underbrace{\int_{\Omega_t} (\widehat{\mathbf{b}}(t, \mathbf{x}) - \mathbf{b}(t, \mathbf{x})) \cdot \nabla q(T-t, \mathbf{x}) d\mathbf{x}.}_{I_2} \end{aligned} \quad (\text{E.1})$$

1058 Recall that $\mathbf{b}(t, \mathbf{x}) = -(\mathbf{f}(T-t, \mathbf{x}) - \frac{1}{2}g(T-t)^2 \nabla \log q_{T-t}(\mathbf{x}))$, $\widehat{\mathbf{b}}(t, \mathbf{x}) = \frac{\partial F}{\partial t}(t, F_t^{-1}(\mathbf{x}))$.
1059

1060 For I_1 , we decompose $\text{tr}[\nabla \widehat{\mathbf{b}}(t, \mathbf{x}) - \nabla \mathbf{b}(t, \mathbf{x})]$ as follows:
1061

$$\begin{aligned} \text{tr}[\nabla \widehat{\mathbf{b}}(t, \mathbf{x}) - \nabla \mathbf{b}(t, \mathbf{x})] &= \left[\nabla \left[\frac{\partial}{\partial t} F \right](t, F_t^{-1}(\mathbf{x})) \nabla F_t^{-1}(\mathbf{x}) + \nabla \mathbf{f}(T-t, \mathbf{x}) \right. \\ &\quad \left. - \frac{1}{2} g(T-t)^2 \nabla^2 \log q_{T-t}(F_t^{-1}(\mathbf{x})) \right] \\ &\quad - \frac{1}{2} g(T-t)^2 \left[\nabla^2 \log q_{T-t}(\mathbf{x}) - \nabla^2 \log q_{T-t}(F_t^{-1}(\mathbf{x})) \right] \\ &= \Psi_k(t, \mathbf{z}) + \left(\nabla \left[\frac{\partial}{\partial t} F \right](t, \mathbf{z}) + \nabla_{\mathbf{z}} [\mathbf{f}(T-t, F_t(\mathbf{z}))] \right) (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \\ &\quad - \frac{1}{2} g(T-t)^2 \left[\nabla^2 \log q_{T-t}(\mathbf{x}) - \nabla^2 \log q_{T-t}(F_t^{-1}(\mathbf{x})) \right], \end{aligned}$$

1072 where $\mathbf{z} = F_t^{-1}(\mathbf{x})$. Therefore, we have
1073

$$\begin{aligned} |I_1| &\leq \int_{\Omega_t} \left| \text{tr} [\nabla \widehat{\mathbf{b}}(t, \mathbf{x}) - \nabla \mathbf{b}(t, \mathbf{x})] \right| q(T-t, \mathbf{x}) d\mathbf{x} \\ &\leq \int |\text{tr} \Psi_k(t, \mathbf{z})| q(T-t, \mathbf{x}) d\mathbf{x} \\ &\quad + \int \left| \text{tr} \left[\nabla \left[\frac{\partial}{\partial t} F \right](t, \mathbf{z}) + \nabla_{\mathbf{z}} [\mathbf{f}(T-t, F_t(\mathbf{z}))] \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right| q(T-t, \mathbf{x}) d\mathbf{x} \end{aligned}$$

$$\begin{aligned}
& + \int \frac{1}{2} g(T-t)^2 |\operatorname{tr} [\nabla^2 \log q_{T-t}(\mathbf{x}) - \nabla^2 \log q_{T-t_k}(\mathbf{z})] | q(T-t, \mathbf{x}) d\mathbf{x} \\
& \stackrel{(i)}{\leq} \left(\int [\operatorname{tr} \Psi_k(t, \mathbf{z})]^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \left(\int \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \\
& \quad + \max_{\mathbf{x}} \left| \operatorname{tr} \left[\nabla \left[\frac{\partial}{\partial t} F \right](t, F_t^{-1}(\mathbf{x})) + \nabla_{\mathbf{z}} [\mathbf{f}(T-t, F_t(\mathbf{z}))] \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right| \\
& \quad + \int \frac{1}{2} g(T-t)^2 |\operatorname{tr} [\nabla^2 \log q_{T-t}(\mathbf{x}) - \nabla^2 \log q_{T-t_k}(\mathbf{z})] | q(T-t, \mathbf{x}) d\mathbf{x} \\
& \stackrel{(ii)}{=} \left(\mathbb{E} [(\operatorname{tr} \Psi_k(t, \mathbf{Y}_{t_k}))^2] \right)^{\frac{1}{2}} \left(\int \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \\
& \quad + \max_{\mathbf{x}} \left| \operatorname{tr} \left[\nabla \left[\frac{\partial}{\partial t} F \right](t, F_t^{-1}(\mathbf{x})) + \nabla_{\mathbf{z}} [\mathbf{f}(T-t, F_t(\mathbf{z}))] \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right| \\
& \quad + \int \frac{1}{2} g(T-t)^2 |\operatorname{tr} [\nabla^2 \log q_{T-t}(\mathbf{x}) - \nabla^2 \log q_{T-t_k}(\mathbf{z})] | q(T-t, \mathbf{x}) d\mathbf{x}. \tag{E.2}
\end{aligned}$$

where (i) holds due to the Cauchy-Schwarz inequality. (ii) holds due to the Jacobian transformation $p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} = p_{\mathbf{Y}_{t_k}}(\mathbf{z}) d\mathbf{z}$ for $\mathbf{z} = F_t^{-1}(\mathbf{x})$.

For I_2 , we decompose $\hat{\mathbf{b}}(t, \mathbf{x}) - \mathbf{b}(t, \mathbf{x})$ as follows:

$$\begin{aligned}
\hat{\mathbf{b}}(t, \mathbf{x}) - \mathbf{b}(t, \mathbf{x}) &= \left[\frac{\partial F}{\partial t} (t, F_t^{-1}(\mathbf{x})) + \mathbf{f}(T-t, \mathbf{x}) - \frac{1}{2} g(T-t)^2 \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x})) \right] \\
&\quad - \frac{1}{2} g(T-t)^2 \left[\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x})) \right] \\
&= \left[\frac{\partial F}{\partial t} (t, \mathbf{z}) + \mathbf{f}(T-t, F_t(\mathbf{z})) - \frac{1}{2} g(T-t)^2 \nabla \log q_{T-t_k}(\mathbf{z}) \right] \\
&\quad - \frac{1}{2} g(T-t)^2 \left[\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x})) \right], \\
&= \Phi_k(t, \mathbf{z}) - \frac{1}{2} g(T-t)^2 \left[\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x})) \right],
\end{aligned}$$

where $\mathbf{z} = F_t^{-1}(\mathbf{x})$. Therefore, we have

$$\begin{aligned}
|I_2| &\leq \int_{\Omega_t} |\Phi_k(t, \mathbf{z}) \cdot \nabla q(T-t, \mathbf{x})| d\mathbf{x} \\
&\quad + \frac{1}{2} g(T-t)^2 \int_{\Omega_t} \left| [\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x}))] \cdot \nabla q(T-t, \mathbf{x}) \right| d\mathbf{x} \\
&\stackrel{(i)}{\leq} \left(\int \|\Phi_k(t, \mathbf{z})\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \left(\int \left\| \frac{\nabla \log q(T-t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \\
&\quad + \frac{1}{2} g(T-t)^2 \int_{\Omega_t} \left| [\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x}))] \cdot \nabla q(T-t, \mathbf{x}) \right| d\mathbf{x} \\
&\stackrel{(ii)}{\leq} \left(\mathbb{E} [\|\Phi_k(t, \mathbf{Y}_{t_k})\|^2] \right)^{\frac{1}{2}} \left(\int \left\| \frac{\nabla \log q(T-t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \right)^{\frac{1}{2}} \\
&\quad + \frac{1}{2} g(T-t)^2 \int_{\Omega_t} \left| [\nabla \log q_{T-t}(\mathbf{x}) - \nabla \log q_{T-t_k}(F_t^{-1}(\mathbf{x}))] \cdot \nabla q(T-t, \mathbf{x}) \right| d\mathbf{x}. \tag{E.3}
\end{aligned}$$

where (i) holds due to the Cauchy-Schwarz inequality. (ii) holds due to the Jacobian transformation $p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} = p_{\mathbf{Y}_{t_k}}(\mathbf{z}) d\mathbf{z}$ for $\mathbf{z} = F_t^{-1}(\mathbf{x})$. Substituting (E.2) and (E.3) into (E.1) and taking the integral over t , we can complete the proof of Theorem 4.3. \square

1134 F ANALYSIS OF VP+EI
 1135

1136 In this section, we focus on the case where $f(\mathbf{x}) = -\mathbf{x}$ and $g(t) = \sqrt{2}$, i.e. the VP-SDE. Specifi-
 1137 cally, the forward process \mathbf{X}_t satisfies:

$$1138 \quad d\mathbf{X}_t = -\mathbf{X}_t dt + \sqrt{2}d\mathbf{W}_t, \quad \mathbf{X}_0 \sim q_0, \quad (\text{F.1})$$

1140 and the true reverse process \mathbf{Y}_t satisfies:

$$1141 \quad d\mathbf{Y}_t = \left(\mathbf{Y}_t + \nabla \log q_{T-t}(\mathbf{Y}_t) \right) dt, \quad \mathbf{Y}_0 \sim q_T.$$

1143 Then, the forward process has the following closed-form expression:
 1144

$$1145 \quad \mathbf{Y}_t = \mathbf{X}_{T-t} = e^{-(T-t)} \mathbf{X}_0 + \sqrt{1 - e^{-2(T-t)}} \mathbf{Z}, \quad \mathbf{Z} \sim N(0, \mathbf{I}_d). \quad (\text{F.2})$$

1146 Recall the exponential integrator provides a discretized version of the reverse process:
 1147

$$1148 \quad \widehat{\mathbf{Y}}_{t_{k+1}} = e^{\eta_k} \widehat{\mathbf{Y}}_{t_k} + \frac{1}{2} (e^{\eta_k} - 1) \frac{1}{2} \mathbf{s}_\theta(T - t_k, \widehat{\mathbf{Y}}_{t_k}).$$

1150 Therefore, we can define the following interpolation operator:
 1151

$$1152 \quad F_t(\mathbf{z}) = e^{t-t_k} \mathbf{z} + (e^{t-t_k} - 1) \mathbf{s}_\theta(T - t_k, \mathbf{z}), \quad (\text{F.3})$$

1153 satisfying $F_{t_{k+1}}(\widehat{\mathbf{Y}}_{t_k}) = \widehat{\mathbf{Y}}_{t_{k+1}}$. Under Assumption 5.4, we know $\mathbf{s}_\theta(t, \cdot)$ is Lipschitz. Therefore,
 1154 suppose $\eta \leq 1/L$, then we have $1/2 \leq \|\nabla F_t\| \leq 3/2$. In particular, F_t is invertible.
 1155 Denote e^{t-t_k} by a . Since $t - t_k \leq \eta \leq 1$, we have $a \leq 1 + 2(t - t_k)$. Then we know that:
 1156

$$\begin{aligned} 1157 \quad \|\mathbf{z}\| &\geq a \|F_t^{-1}(\mathbf{z})\| - (a - 1) \left(L \|F_t^{-1}(\mathbf{z})\| + c \right) \\ 1158 \quad &\geq \|F_t^{-1}(\mathbf{z})\| - 2(t - t_k) \left(L \|F_t^{-1}(\mathbf{z})\| + c \right) \\ 1159 \quad &\geq \frac{1}{2} \|F_t^{-1}(\mathbf{z})\| - \frac{1}{2}, \\ 1160 \quad & \\ 1161 \quad & \\ 1162 \quad & \\ 1163 \quad & \end{aligned}$$

where the last inequality holds when we assume that $t - t_k \leq \eta \leq \min\{1/4L, 1/4c\}$. Thus we
 1164 know that:

$$1165 \quad \|F_t^{-1}(\mathbf{z})\| \leq 2\|\mathbf{z}\| + 1. \quad (\text{F.4})$$

1166 Before we start the proof, we introduce several important lemmas used in our proof. Let $c_1 =$
 1167 $(1 + d/2)4L^2$ and $c_2 = (1 + d/2)(L + c)^2$. Using our assumption on η , we know that
 1168

$$1169 \quad \eta \leq \min \left\{ \frac{1}{16d}, \frac{1}{2c_1d}, \frac{1}{c_2}, \frac{1}{4R^2c_1} \right\}. \quad (\text{F.5})$$

1170 The following two lemmas are related to the ratios $p_{\mathbf{Y}_t}(\mathbf{x})/p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})$.

1171 **Lemma F.1.** Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (F.2) and (F.3). Under Assumptions 5.4 and 5.1,
 1172 suppose our time schedule satisfies (2.6). Then, for $\eta \leq \min\{\frac{1}{2(L+1)d}, \frac{1}{16d}, \frac{1}{c_1d}, \frac{1}{2c_1}, \frac{1}{c_2}, \frac{1}{R^2c_1}\}$, we
 1173 have

$$1174 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2}.$$

1175 Moreover, we have

$$1176 \quad \int_{\Omega_t} \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim 1.$$

1177 **Lemma F.2.** Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (F.2) and (F.3). Under Assumptions 5.1 and 5.4,
 1178 suppose our time schedule satisfies (2.6). Then for $\eta \leq \min\{\frac{1}{2(L+1)d}, \frac{1}{16d}, \frac{1}{2c_1d}, \frac{1}{2c_1}, \frac{1}{c_2}, \frac{1}{4R^2c_1}\}$,
 1179 we have:

$$1180 \quad \int \left\| \frac{\nabla q(T - t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim \frac{d}{\min\{T - t, 1\}}.$$

The following two lemmas show the upper bound of the derivatives of the true score function with respect to t and \mathbf{x} separately.

Lemma F.3. When \mathbf{X}_t is defined as in (F.1), under Assumption 5.1, we have:

$$\begin{aligned} \left| \frac{\partial}{\partial t} \left(\text{tr} (\nabla^2 \log q(t, \mathbf{x})) \right) \right| &\lesssim \frac{d}{\min\{t, 1\}^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{\min\{t, 1\}^4}, \\ \left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| &\lesssim \frac{\|\mathbf{x}\| + R}{\min\{t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^2 R^2}{\min\{t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^3}{\min\{t, 1\}^3}. \end{aligned}$$

Lemma F.4. When \mathbf{X}_t is defined as in (F.1), under Assumption 5.1, we have:

$$\begin{aligned} \|\nabla \log q(t, \mathbf{x})\| &\lesssim \frac{\|\mathbf{x}\| + R}{\min\{t, 1\}}, \\ \|\nabla^2 \log q(t, \mathbf{x})\|_2 &\leq \frac{1}{\min\{t, 1\}} + \frac{R^2}{\min\{t, 1\}^2}, \\ \|\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x}))\| &\leq \frac{R^3}{\min\{t, 1\}^3}. \end{aligned}$$

The proofs of the above lemmas can be found in later sections.

F.1 PROOF OF THEOREM 5.5

Proof of Theorem 5.5. From the discussion in Section 5.1, we know that:

$$\Phi_k(t, \mathbf{x}) = s_\theta(T - t_k, \mathbf{x}) - \nabla \log q_{T-t_k}(\mathbf{x}), \quad (\text{F.6})$$

$$\Psi_k(t, \mathbf{x}) = \nabla s_\theta(T - t_k, \mathbf{x}) - \nabla^2 \log q_{T-t_k}(\mathbf{x}), \quad (\text{F.7})$$

in this specific case. Recall that $g(t) = \sqrt{2}$ and $\mathbf{f}(T - t, \mathbf{x}) = -\mathbf{x} = -F_t(\mathbf{z})$. From (F.3) we have:

$$\frac{\partial}{\partial t} F(t, \mathbf{x}) = F_t(\mathbf{x}) + s_\theta(T - t_k, \mathbf{x}).$$

Thus we know that

$$\nabla \left[\frac{\partial}{\partial t} F \right] (t, F_t^{-1}(\mathbf{x})) = [\nabla F_t](\mathbf{z}) + \nabla s_\theta(T - t_k, \mathbf{z}). \quad (\text{F.8})$$

By Theorem 4.3, we know that:

$$\begin{aligned} \text{TV}(\mathbf{Y}_{T-\delta}, \widehat{\mathbf{Y}}_{t_N}) &\leq \text{TV}(q_T, \pi_d) \\ &+ \sum_{k=0}^{N-1} \int_{t_k}^{t_{k+1}} \underbrace{\left[\sqrt{\mathbb{E}[\|\Phi_k(t, \mathbf{Y}_{t_k})\|^2]} \right]}_{J_1: \text{Score estimation error}} \sqrt{\int \left\| \frac{\nabla \log q(T - t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}} \\ &+ \underbrace{\sqrt{\mathbb{E}[\text{tr}(\Psi_k(t, \mathbf{Y}_{t_k}))^2]} \sqrt{\int \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}}}_{J_2: \text{Divergence estimation error}} \\ &+ \underbrace{\int |(\nabla \log q_{T-t_k}(\mathbf{z}) - \nabla \log q_{T-t}(\mathbf{x})) \cdot \nabla \log q(T - t, \mathbf{x})| q(T - t, \mathbf{x}) d\mathbf{x}}_{J_3: \text{Score discretization error}} \\ &+ \underbrace{\int |\text{tr}(\nabla^2 \log q(T - t_k, \mathbf{z}) - \nabla^2 \log q(T - t, \mathbf{x}))| q(T - t, \mathbf{x}) d\mathbf{x}}_{J_4: \text{Divergence discretization error}} \\ &+ \underbrace{\max_{\mathbf{x}} \left| \text{tr} \left[\nabla s_\theta(T - t_k, \mathbf{z}) \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right| \right] dt. \end{aligned} \quad (\text{F.9})$$

Bounding the Score Estimation Error J_1 . We can easily verify that our η is small enough to satisfies the condition of Lemma F.2. Therefore, using Lemma F.2 and (F.6), we have=

$$|J_1| \lesssim \sqrt{\frac{d}{\min\{T-t, 1\}}} \sqrt{\mathbb{E}_Q \left\| \mathbf{s}_\theta(T-t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2}. \quad (\text{F.10})$$

Bounding the Divergence Estimation Error J_2 .

Similarly, our η is small enough to satisfies the condition of Lemma F.1. Using Lemma F.1 and (F.7), we have

$$|J_2| \lesssim \sqrt{\mathbb{E}_Q \operatorname{tr} \left(\nabla \mathbf{s}_\theta(T-t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q(T-t_k, \mathbf{Y}_{t_k}) \right)^2}. \quad (\text{F.11})$$

Bounding the Bias Term J_5 . From (F.3), we know that:

$$\begin{aligned} \nabla F_t^{-1}(\mathbf{x}) &= [(\nabla F_t)(F_t^{-1}(\mathbf{x}))]^{-1} \\ &= [\mathbf{e}^{t-t_k} \mathbf{I}_d + (\mathbf{e}^{t-t_k} - 1) \nabla \mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x}))]^{-1}. \end{aligned}$$

Using Assumption 5.4 and $t-t_k \leq \eta \leq 1$, we have:

$$\begin{aligned} \|(\nabla F_t)(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 &\leq (\mathbf{e}^{t-t_k} - 1) \|\mathbf{I}_d + \nabla \mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\|_2 \\ &\leq 2(t-t_k)(1+L). \end{aligned}$$

Let $\mathbf{A} = (\nabla F_t)(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d$. Then we have $\|\mathbf{A}\|_2 \leq 2(t-t_k)(1+L)$. Thus we know that:

$$\begin{aligned} \|\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 &= \|(\mathbf{I}_d + \mathbf{A})^{-1} - \mathbf{I}_d\|_2 \\ &\stackrel{(i)}{=} \left\| \sum_{n=1}^{\infty} (-1)^n \mathbf{A}^n \right\|_2 \\ &\stackrel{(ii)}{\leq} \sum_{n=1}^{\infty} \|\mathbf{A}\|_2^n \\ &\stackrel{(iii)}{\leq} 4(t-t_k)(1+L), \end{aligned} \quad (\text{F.12})$$

here (i) holds because of the series expansion for $\|\mathbf{A}\|_2 < 1$. (ii) holds due to the Cauchy-Schwarz inequality. (iii) holds due to our assumption $\eta \leq 1/4(1+L)$. Then we know that:

$$\begin{aligned} |J_5| &= \max_{\mathbf{x}} \left| \operatorname{tr} \left[\nabla \mathbf{s}_\theta(T-t_k, \mathbf{z}) (\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d) \right] \right| \\ &\stackrel{(i)}{\leq} d \max \left\| \nabla \mathbf{s}_\theta(T-t_k, \mathbf{z}) (\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d) \right\|_2 \\ &\leq d \max \|\nabla \mathbf{s}_\theta(T-t_k, \mathbf{z})\|_2 \|\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 \\ &\stackrel{(ii)}{\leq} 4dL(t-t_k)(1+L), \end{aligned}$$

where (i) holds due to $\operatorname{tr}(\mathbf{A}) \leq d\|\mathbf{A}\|_2$. (ii) holds due to Assumption 5.4 and (F.12). Thus we have:

$$|J_5| \lesssim (t-t_k)dL^2. \quad (\text{F.13})$$

Bounding the Moments of \mathbf{Y}_t .

Before we start estimating J_3 and J_4 , we first do some preparation. Since we have $\mathbf{Y}_t = \mathbf{X}_{T-t} = \mathbf{e}^{-(T-t)} \mathbf{X}_0 + \sqrt{1 - \mathbf{e}^{-2(T-t)}} \mathbf{Z}$, here $\mathbf{Z} \sim N(0, \mathbf{I}_d)$. Our goal is to bound the moments of \mathbf{Y}_t . When $T-t < 1$, we know that $\mathbf{e}^{-(T-t)} = \Theta(1)$ and $\sqrt{1 - \mathbf{e}^{-2(T-t)}} = O(\sqrt{T-t})$. Thus we have:

$$\begin{aligned} \mathbb{E}\|\mathbf{Y}_t\|^2 &= \mathbf{e}^{-2(T-t)} \mathbb{E}\|\mathbf{X}_0\|^2 + (1 - \mathbf{e}^{-2(T-t)}) \mathbb{E}\|\mathbf{Z}\|^2 + \mathbb{E}\mathbf{e}^{-(T-t)} \sqrt{1 - \mathbf{e}^{-2(T-t)}} \langle \mathbf{X}_0, \mathbf{Z} \rangle \\ &\lesssim \mathbb{E}\|\mathbf{X}_0\|^2 + (T-t)\mathbb{E}\|\mathbf{Z}\|^2 \\ &= R^2 + (T-t)d. \end{aligned} \quad (\text{F.14})$$

where the first inequality holds due to the independence of \mathbf{X}_0 and \mathbf{Z} . When $T-t \geq 1$, we have:

$$\mathbb{E}\|\mathbf{Y}_t\|^2 \lesssim \mathbf{e}^{-2(T-t)} \mathbb{E}\|\mathbf{X}_0\|^2 + \mathbb{E}\|\mathbf{Z}\|^2$$

$$= e^{-2(T-t)}R^2 + d. \quad (\text{F.15})$$

Combine (F.14) and (F.15) together, we know that:

$$\mathbb{E}\|\mathbf{Y}_t\|^2 \lesssim R^2 + \min\{T-t, 1\}d. \quad (\text{F.16})$$

After computing the second-order moment, using the inequality $(\mathbb{E}\|\mathbf{Y}_t\|)^2 \leq \mathbb{E}\|\mathbf{Y}_t\|^2$, we can bound the first-order moment as follows:

$$\mathbb{E}\|\mathbf{Y}_t\| \lesssim R + \sqrt{\min\{T-t, 1\}}\sqrt{d}. \quad (\text{F.17})$$

Moreover, we consider the fourth-order moment using similar methods. When $T-t < 1$, we have $e^{-(T-t)} = \Theta(1)$ and $\sqrt{1-e^{-(T-t)}} = O(\sqrt{T-t})$. Thus we know that:

$$\begin{aligned} \mathbb{E}\|\mathbf{Y}_t\|^4 &\lesssim \mathbb{E}\|\mathbf{X}_0\|^4 + (T-t)^2\mathbb{E}\|\mathbf{Z}\|^4 \\ &\lesssim R^4 + (T-t)^2d^2, \end{aligned} \quad (\text{F.18})$$

where the first inequality holds due to $\|\mathbf{x} - \mathbf{y}\|^4 \leq (\|\mathbf{x}\| + \|\mathbf{y}\|)^4 \leq C(\|\mathbf{x}\|^4 + \|\mathbf{y}\|^4)$ for some constant C . Same as what we did earlier, when $T-t \geq 1$, we have:

$$\begin{aligned} \mathbb{E}\|\mathbf{Y}_t\|^4 &\lesssim e^{-4(T-t)}\mathbb{E}\|\mathbf{X}_0\|^4 + \mathbb{E}\|\mathbf{Z}\|^4 \\ &\lesssim e^{-3(T-t)}R^4 + d^2. \end{aligned} \quad (\text{F.19})$$

Putting (F.18) and (F.19) together, we have:

$$\mathbb{E}\|\mathbf{Y}_t\|^4 \lesssim R^4 + \min\{T-t, 1\}^2d^2. \quad (\text{F.20})$$

After computing the fourth-order momentum, using the inequality $\mathbb{E}\|\mathbf{Y}_t\|^3 \leq \sqrt{\mathbb{E}\|\mathbf{Y}_t\|^2} \cdot \sqrt{\mathbb{E}\|\mathbf{Y}_t\|^4}$, we know that:

$$\mathbb{E}\|\mathbf{Y}_t\|^3 \lesssim R^3 + \min\{T-t, 1\}^{3/2}d^{3/2}. \quad (\text{F.21})$$

Bounding Divergence Discretization Error J_4 . Let $\mathbf{z} = F_t^{-1}(\mathbf{x})$. We start by bounding $\|\mathbf{z} - \mathbf{x}\|$. We know that

$$\begin{aligned} \|\mathbf{z} - \mathbf{x}\| &= \|F_t^{-1}(\mathbf{x}) - \mathbf{x}\| \\ &= (e^{t-t_k} - 1)\|F_t^{-1}(\mathbf{x}) + s_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\| \\ &\leq 2(t-t_k)\|F_t^{-1}(\mathbf{x}) + s_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\| \\ &\leq 2(t-t_k)(L+1)\|F_t^{-1}(\mathbf{x})\| + (t-t_k)c, \end{aligned}$$

where the first inequality holds due to $t-t_k < 1$. The second inequality holds due to Assumption 5.4. From (F.4) we know that $\|F_t^{-1}(\mathbf{x})\| \leq 2\|\mathbf{x}\| + 1$. Hence we have:

$$\begin{aligned} \|\mathbf{z} - \mathbf{x}\| &\lesssim (t-t_k)(L+1)\|\mathbf{x}\| + (t-t_k)(L+c+1) \\ &\stackrel{(i)}{\lesssim} (t-t_k)L(\|\mathbf{x}\| + 1). \end{aligned} \quad (\text{F.22})$$

Using Lemma F.3 and Lemma F.4, we have

$$\begin{aligned} &\left| \int_{\Omega_t} \text{tr} \left(\nabla^2 \log q(T-t_k, \mathbf{z}) - \nabla^2 \log q(T-t, \mathbf{x}) \right) q(T-t, \mathbf{x}) d\mathbf{x} \right| \\ &\lesssim \int_{\Omega_t} \left((t-t_k) \left(\frac{d}{\min\{T-t, 1\}^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{\min\{T-t, 1\}^4} \right) + \frac{R^3}{\min\{T-t, 1\}^3} \|\mathbf{x} - \mathbf{z}\| \right) q(T-t, \mathbf{x}) d\mathbf{x} \\ &\stackrel{(i)}{\lesssim} \int_{\Omega_t} \left((t-t_k) \left(\frac{d}{\min\{T-t, 1\}^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{\min\{T-t, 1\}^4} \right) \right. \\ &\quad \left. + \frac{R^3}{\min\{T-t, 1\}^3} (t-t_k)L(\|\mathbf{x}\| + 1) \right) q(T-t, \mathbf{x}) d\mathbf{x} \\ &\lesssim (t-t_k) \frac{d}{\min\{T-t, 1\}^2} + (t-t_k) \frac{R^2}{\min\{T-t, 1\}^4} \mathbb{E}(\|\mathbf{Y}_t\|^2 + R^2) \end{aligned}$$

$$\begin{aligned}
& + (t - t_k) \frac{R^3}{\min\{T - t, 1\}^3} L \cdot \mathbb{E}(\|\mathbf{Y}_t\| + 1) \\
& \stackrel{(ii)}{\lesssim} (t - t_k) \frac{d}{\min\{T - t, 1\}^2} + (t - t_k) \frac{R^4 + R^2 \min\{T - t, 1\} d}{\min\{T - t, 1\}^4} \\
& \quad + (t - t_k) \frac{LR^3}{\min\{T - t, 1\}^3} (R + \sqrt{T - t} \sqrt{d} + 1),
\end{aligned}$$

where (i) holds due to (F.22). (ii) holds due to (F.16) and (F.17). Assuming $R \geq 1$, we can further simplify its form:

$$\begin{aligned}
|J_4| & \lesssim (t - t_k) d \frac{R^2}{\min\{T - t, 1\}^3} + (t - t_k) (1 + L \min\{T - t, 1\}) \frac{R^4}{\min\{T - t, 1\}^4} \\
& \quad + (t - t_k) \sqrt{d} \frac{LR^3}{\min\{T - t, 1\}^{5/2}} \\
& \lesssim (t - t_k) d \frac{R^2}{\min\{T - t, 1\}^3} + (t - t_k) \frac{LR^4}{\min\{T - t, 1\}^4} + (t - t_k) \sqrt{d} \frac{LR^3}{\min\{T - t, 1\}^{5/2}}.
\end{aligned} \tag{F.23}$$

Bounding Score Discretization Error J_3 .

According to Lemma F.4, we know that:

$$\begin{aligned}
\|\nabla \log q(T - t, \mathbf{x})\| & \lesssim \frac{\|\mathbf{x}\| + R}{\min\{T - t, 1\}}, \\
\|\nabla^2 \log q(T - t, \mathbf{x})\|_2 & \lesssim \frac{1}{\min\{T - t, 1\}} + \frac{R^2}{\min\{T - t, 1\}^2} \stackrel{(i)}{\lesssim} \frac{R^2}{\min\{T - t, 1\}^2}.
\end{aligned}$$

Here (i) is because we assumed that $R \geq 1$. Using Lemma F.3, we have:

$$\begin{aligned}
\left\| \frac{\partial}{\partial t} (\nabla \log q(T - t, \mathbf{x})) \right\| & \lesssim \frac{\|\mathbf{x}\| + R}{\min\{T - t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^2 R^2}{\min\{T - t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^3}{\min\{T - t, 1\}^3} \\
& \lesssim \frac{(\|\mathbf{x}\| + R)^2 R^2}{\min\{T - t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^3}{\min\{T - t, 1\}^3}.
\end{aligned}$$

Thus we know that:

$$\begin{aligned}
& \left| \int_{\Omega_t} (\nabla \log q_{T-t_k}(\mathbf{z}) - \nabla \log q_{T-t}(\mathbf{x})) \nabla \log q(T - t, \mathbf{x}) q(T - t, \mathbf{x}) d\mathbf{x} \right| \\
& \lesssim \int_{\Omega_t} (t - t_k) \left(\frac{(\|\mathbf{x}\| + R)^2 R^2}{\min\{T - t, 1\}^2} + \frac{(\|\mathbf{x}\| + R)^3}{\min\{T - t, 1\}^3} \right) \frac{\|\mathbf{x}\| + R}{\min\{T - t, 1\}} q(T - t, \mathbf{x}) d\mathbf{x} \\
& \quad + \int_{\Omega_t} \frac{R^2}{\min\{T - t, 1\}^2} \|\mathbf{x} - \mathbf{z}\| \frac{\|\mathbf{x}\| + R}{\min\{T - t, 1\}} q(T - t, \mathbf{x}) d\mathbf{x} \\
& \lesssim (t - t_k) \frac{R^2}{\min\{T - t, 1\}^3} \mathbb{E}(\|\mathbf{Y}_t\| + R)^3 + (t - t_k) \frac{1}{\min\{T - t, 1\}^4} \mathbb{E}(\|\mathbf{Y}_t\| + R)^4 \\
& \quad + (t - t_k) \int_{\Omega_t} \frac{R^2}{\min\{T - t, 1\}^2} L(\|\mathbf{x}\| + 1) \frac{\|\mathbf{x}\| + R}{\min\{T - t, 1\}} q(T - t, \mathbf{x}) d\mathbf{x} \\
& \stackrel{(i)}{\lesssim} (t - t_k) \frac{R^2}{\min\{T - t, 1\}^3} (R^3 + \min\{T - t, 1\}^{3/2} d^{3/2}) \\
& \quad + (t - t_k) \frac{1}{\min\{T - t, 1\}^4} (R^4 + \min\{T - t, 1\}^2 d^2) \\
& \quad + (t - t_k) \frac{LR^2}{\min\{T - t, 1\}^3} (R^2 + \min\{T - t, 1\} d),
\end{aligned} \tag{F.24}$$

where we use (F.20) and (F.21) to bound the third and fourth moments of \mathbf{Y}_t . (i) holds due to:

$$\int_{\Omega_t} (\|\mathbf{x}\| + 1)(\|\mathbf{x}\| + R) q(T - t, \mathbf{x}) d\mathbf{x} \lesssim \mathbb{E}(\|\mathbf{Y}_t\| + R)^2 \lesssim R^2 + \min\{T - t, 1\} d.$$

1404 Reorganizing terms in (F.24), we have:
 1405

$$\begin{aligned} |J_3| &\lesssim (t - t_k) \left[\frac{R^5}{\min\{T-t, 1\}^3} + d^{3/2} \frac{R^2}{\min\{T-t, 1\}^{3/2}} + \frac{R^4}{\min\{T-t, 1\}^4} \right. \\ &\quad \left. + d^2 \frac{1}{\min\{T-t, 1\}^2} + \frac{LR^4}{\min\{T-t, 1\}^3} + d \frac{LR^2}{\min\{T-t, 1\}^2} \right]. \end{aligned} \quad (\text{F.25})$$

1411 **Putting J_3 and J_4 Together.**

1412 By (F.23) and (F.25), and since we assume $R \geq 1$ and $L \geq 1$, we have:

$$\begin{aligned} |J_3| + |J_4| &\lesssim (t - t_k) \underbrace{d \frac{R^2}{\min\{T-t, 1\}^3}}_{C_1} + (t - t_k) \underbrace{\frac{LR^4}{\min\{T-t, 1\}^4}}_{C_2} + (t - t_k) \underbrace{\sqrt{d} \frac{LR^3}{\min\{T-t, 1\}^{5/2}}}_{C_3} \\ &\quad + (t - t_k) \left[\underbrace{\frac{R^5}{\min\{T-t, 1\}^3}}_{C_4} + d^{3/2} \underbrace{\frac{R^2}{\min\{T-t, 1\}^{3/2}}}_{C_5} + \underbrace{\frac{R^4}{\min\{T-t, 1\}^4}}_{C_6} \right. \\ &\quad \left. + d^2 \underbrace{\frac{1}{\min\{T-t, 1\}^2}}_{C_7} + \underbrace{\frac{LR^4}{\min\{T-t, 1\}^3}}_{C_8} + d \underbrace{\frac{LR^2}{\min\{T-t, 1\}^2}}_{C_9} \right] \\ &\stackrel{(ii)}{\lesssim} (t - t_k) [C_1 + C_2 + C_3 + C_4 + C_5 + C_7 + C_9] \\ &= (t - t_k) \left[d^2 \frac{1}{\min\{T-t, 1\}^2} + d^{3/2} \frac{R^2}{\min\{T-t, 1\}^{3/2}} + d \left(\frac{LR^2 \min\{T-t, 1\} + R^2}{\min\{T-t, 1\}^3} \right) \right. \\ &\quad \left. + \sqrt{d} \frac{LR^3}{\min\{T-t, 1\}^{5/2}} + \frac{R^5 \min\{T-t, 1\} + LR^4}{\min\{T-t, 1\}^4} \right]. \end{aligned} \quad (\text{F.26})$$

1431 Here (ii) is because $C_6 \leq C_2$ and $C_8 \leq C_2$.
 1432

1433 **Combining Everything Together.**

1434 Plugging (F.10) (F.11) (F.13) (F.26) back to (F.9), and taking the integral, we can obtain the following
 1435 inequality:

$$\begin{aligned} \text{TV}(\hat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) &\lesssim \text{TV}(\pi_d, q_T) \\ &\quad + \underbrace{\sum_k (t_{k+1} - t_k) \sqrt{\mathbb{E}_Q \text{tr} \left(\nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q(T - t_k, \mathbf{Y}_{t_k}) \right)^2}}_{K_1} \\ &\quad + \underbrace{\sum_k (t_{k+1} - t_k) \sqrt{\frac{d}{\min\{T - t_{k+1}, 1\}}} \sqrt{\mathbb{E}_Q \left\| \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2}}_{K_2} \\ &\quad + \underbrace{\sum_k (t_{k+1} - t_k)^2 d L^2}_{K_3} \\ &\quad + \sum_k (t_{k+1} - t_k)^2 \left[d^2 \frac{1}{\min\{T-t, 1\}^2} + d^{3/2} \frac{R^2}{\min\{T-t, 1\}^{3/2}} + d \left(\frac{LR^2 \min\{T-t, 1\} + R^2}{\min\{T-t, 1\}^3} \right) \right. \\ &\quad \left. + \sqrt{d} \frac{LR^3}{\min\{T-t, 1\}^{5/2}} + \frac{R^5 \min\{T-t, 1\} + LR^4}{\min\{T-t, 1\}^4} \right]. \end{aligned}$$

1454 Using Assumption 5.3, we have $K_1 = \epsilon_{\text{div}}$. For K_2 , using the Cauchy-Schwartz inequality, we have
 1455

$$\sum_k (t_{k+1} - t_k) \sqrt{\frac{d}{\min\{T - t_{k+1}, 1\}}} \sqrt{\mathbb{E}_Q \left\| \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2}$$

$$\begin{aligned}
&\leq \sqrt{\sum_k (t_{k+1} - t_k) \frac{d}{\min\{T - t_{k+1}, 1\}}} \sqrt{\sum_k (t_{k+1} - t_k) \mathbb{E}_Q \left\| \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2} \\
&\leq \sqrt{\eta N d} \epsilon_{\text{score}}.
\end{aligned}$$

For K_3 , we know that:

$$\sum_k (t_{k+1} - t_k)^2 dL^2 \leq \eta^2 N d L^2.$$

For K_4 , using our assumption on time schedule $t_{k+1} - t_k \leq \eta (T - t_{k+1})$, we have:

$$\begin{aligned}
K_4 &\lesssim \eta^2 N \left[d^2 + d^{3/2} R^2 + LR^2 d + R^2 d \frac{1}{\delta} \right. \\
&\quad \left. + LR^3 \sqrt{\frac{d}{\delta}} + R^5 \frac{1}{\delta} + LR^4 \frac{1}{\delta^2} \right] \\
&\stackrel{(ii)}{\lesssim} \eta^2 N \left[LR^4 d^2 + d^{3/2} R^2 + LR^2 d + R^5 \frac{1}{\delta} + LR^4 \frac{1}{\delta^2} \right] \\
&\stackrel{(iii)}{\lesssim} \eta^2 N \left[LR^4 d^2 + R^5 \frac{1}{\delta} + LR^4 \frac{1}{\delta^2} \right],
\end{aligned}$$

where (ii) holds due to $R^2 d \frac{1}{\delta} \leq d^2 + \frac{LR^4}{\delta^2}$ and $2LR^3 \sqrt{\frac{d}{\delta}} \leq LR^2 d + LR^4 \frac{1}{\delta^2}$. (iii) holds due to $2d^{3/2} R^2 \leq LR^4 d^2 + LR^2 d$ and $2LR^2 d \leq LR^4 d^2 + LR^4 \frac{1}{\delta^2}$. Putting K_1, K_2, K_3 and K_4 together, we have

$$\begin{aligned}
\text{TV}(\hat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) &\lesssim \text{TV}(q_T, \pi_d) + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} \\
&\quad + \eta^2 N \left[LR^4 \left(d^2 + \frac{1}{\delta^2} \right) + L^2 d + \frac{R^5}{\delta} \right].
\end{aligned}$$

This completes the proof of Theorem 5.5. \square

F.2 PROOF OF LEMMA 5.6 AND COROLLARY 5.7

Proof of Lemma 5.6. By proposition 4 in Benton et al. (2024), we have:

$$\text{KL}(q_T \| \pi_d) \lesssim d e^{-2T}, \text{ for } T \geq 1.$$

Then by Pinsker's inequality, we know that:

$$\text{TV}(q_T, \pi_d) \leq \sqrt{\frac{1}{2} \text{KL}(q_T \| \pi_d)} \lesssim \sqrt{d} e^{-T}.$$

This completes the proof. \square

Proof of Corollary 5.7. Please refer to Benton et al. (2024)'s Appendix D for a detailed derivation of the existence of such time schedule. Since we take $T = \log(d/\epsilon^2)/2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$, using $\epsilon \leq 1/L$, we can easily show that $\eta \leq \min\{1/(12L^2 d^2), 1/(24L^2 R^2 d)\}$. Hence by Theorem 5.5 and Lemma 5.6, we have:

$$\text{TV}(\hat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \sqrt{d} e^{-T} + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N \left[LR^4 \left(d^2 + \frac{1}{\delta^2} \right) + L^2 d + \frac{R^5}{\delta} \right].$$

Since we take $\delta = 1/d$ and assumed that $d \geq R/L + L/R^4$, we can further simply the bound of TV to:

$$\text{TV}(\hat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \sqrt{d} e^{-T} + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N L R^4 d^2.$$

Since $T = \log(d/\epsilon^2)/2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$, we know that

$$\begin{aligned}
\sqrt{d} e^{-T} &\leq \epsilon, \\
\eta^2 N L R^4 d^2 &= \frac{1}{N} (T + \log(1/\delta))^2 L R^4 d^2 \leq \epsilon.
\end{aligned}$$

Recall that we assume sufficiently small score estimation and divergence error, this completes the proof of Corollary 5.7. \square

1512 **G ANALYSIS OF VE+DDIM**
 1513

1514 In this section, we focus on the case where $f(\mathbf{x}) = 0$ and $g(t) = 1$, i.e. taking $\sigma(t)^2 = t$ in VE-SDE.
 1515 Specifically, the forward process \mathbf{X}_t satisfies:

$$1516 \quad d\mathbf{X}_t = d\mathbf{W}_t, \quad \mathbf{X}_0 \sim q_0, \quad (\text{G.1})$$

1517 and the true reverse process \mathbf{Y}_t satisfies:

$$1519 \quad d\mathbf{Y}_t = \frac{1}{2} \nabla \log q_{T-t}(\mathbf{Y}_t) dt, \quad \mathbf{Y}_0 \sim q_T. \quad (\text{G.2})$$

1521 Then, the forward process has the following closed-form expression:

$$1522 \quad \mathbf{Y}_t = \mathbf{X}_{T-t} = \mathbf{X}_0 + \sqrt{T-t}\mathbf{Z}, \quad \mathbf{Z} \sim N(0, \mathbf{I}_d). \quad (\text{G.3})$$

1524 Recall that we interpret DDIM as a numerical scheme, and it provides the following discretized
 1525 version of the reverse process:

$$1526 \quad \widehat{\mathbf{Y}}_{t_{k+1}} = \widehat{\mathbf{Y}}_{t_k} + (t_{k+1} - t_k)l\left(1 - \sqrt{1 - \frac{1}{l}}\right)\mathbf{s}_\theta(T - t_k, \widehat{\mathbf{Y}}_{t_k}), \quad l = \frac{T - t_k}{t_{k+1} - t_k}. \quad (\text{G.4})$$

1529 Therefore, denoting $l(1 - \sqrt{1 - 1/l})$ by c_l , we have $c_l \leq 1$. Then we can define the following
 1530 interpolation operator:

$$1531 \quad F_t(\mathbf{z}) = \mathbf{z} + (t - t_k)c_l\mathbf{s}_\theta(T - t_k, \mathbf{z}), \quad (\text{G.5})$$

1533 satisfying $F_{t_{k+1}}(\widehat{\mathbf{Y}}_{t_k}) = \widehat{\mathbf{Y}}_{t_{k+1}}$. Under Assumption 5.4, we know $\mathbf{s}_\theta(t, \cdot)$ is Lipschitz. Therefore,
 1534 suppose $\eta \leq 1/L$, then we have $1/2 \leq \|\nabla F_t\| \leq 3/2$. In particular, F_t is invertible.

1535 Moreover, using Assumption 5.4 and $F_t^{-1}(\mathbf{x}) - \mathbf{x} = -(t - t_k)c_l \cdot \mathbf{s}_\theta(T - t_k, F_t^{-1}(\mathbf{x}))$, when
 1536 $\eta \leq \frac{1}{L+c}$, we have

$$1538 \quad \|F_t^{-1}(\mathbf{x}) - \mathbf{x}\| = (t - t_k)c_l\|\mathbf{s}_\theta(T - t_k, F_t^{-1}(\mathbf{x}))\| \\ 1539 \quad \leq (t - t_k)L\|F_t^{-1}(\mathbf{x})\| + (t - t_k)c_l \\ 1540 \quad \leq \|\mathbf{x}\| + 1. \quad (\text{G.6})$$

1541 Thus we know that

$$1543 \quad \|F_t^{-1}(\mathbf{x})\| \leq 2\|\mathbf{x}\| + 1. \quad (\text{G.7})$$

1544 Before we start the proof, we introduce several important lemmas used in our proof. Let $c_1 = (1 + d/2)4L^2$ and $c_2 = (1 + d/2)(L + c)^2$. Using our assumption on η , we know that

$$1547 \quad \eta \leq \min\left\{\frac{1}{4c_1dT}, \frac{1}{c_2}, \frac{1}{4R^2c_1}\right\}. \quad (\text{G.8})$$

1549 The following two lemmas are related to the ratios $p_{\mathbf{Y}_t}(\mathbf{x})/p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})$.

1551 **Lemma G.1.** Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (G.2) and (G.5). Under Assumption 5.4 (Lipschitz
 1552 condition on \mathbf{s}_θ), Assumption 5.1 (bounded support) and suppose the time schedule satisfies (2.6),
 1553 assuming $\eta \leq \min\{\frac{1}{d}, \frac{1}{L+c}, \frac{1}{4c_1(T-t)}, \frac{1}{c_1R^2}, \frac{1}{c_2}\}$, we have

$$1555 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2},$$

1557 where $c_1 = (1 + d/2)4L^2$ and $c_2 = (1 + d/2)(L + c)^2$. Moreover, we have

$$1559 \quad \int_{\Omega_t} \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim 1.$$

1561 **Lemma G.2.** Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (G.2) and (G.5). Under Assumption 5.4, 5.1 and 2.6,
 1562 suppose $\eta \leq \min\{\frac{1}{c_2}, \frac{1}{c_1R^2}, \frac{1}{2(d+2)c_1(T-t)}\}$, where $c_1 = (1 + d/2)4L^2$ and $c_2 = (1 + d/2)(L + c)^2$.
 1563 then, we have:

$$1565 \quad \int_{\Omega_t} \left\| \frac{\nabla q(T-t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim \frac{d}{T-t}.$$

The following two lemmas show the upper bound of the derivatives of the true score function of the VE forward process concerning t and \mathbf{x} separately.

Lemma G.3. When \mathbf{X}_t is defined as in (G.1), then under Assumption 5.1, we have:

$$\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| \lesssim \frac{\|\mathbf{x}\| + R}{t^2} + \frac{(\|\mathbf{x}\| + R)^3}{t^3}.$$

Moreover, we have:

$$\left| \frac{\partial}{\partial t} (\text{tr}(\nabla^2 \log q(t, \mathbf{x}))) \right| \lesssim \frac{d}{t^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4}.$$

Lemma G.4. When \mathbf{X}_t is defined as in (G.1), then under Assumption 5.1, we have:

$$\begin{aligned} \|\nabla \log q(t, \mathbf{x})\| &\lesssim \frac{\|\mathbf{x}\| + R}{t}, \\ \|\nabla^2 \log q(t, \mathbf{x})\|_2 &\leq \frac{1}{t} + \frac{R^2}{t^2}, \\ \|\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x}))\| &\leq \frac{R^3}{t^3}. \end{aligned}$$

The proofs of the above lemmas can be found in later sections.

G.1 PROOF OF THEOREM 5.9

Proof of Theorem 5.9. From the discussion in Section 5.2, we know that:

$$\Phi_k(t, \mathbf{x}) = c_l \mathbf{s}_\theta(T - t_k, \mathbf{x}) - \frac{1}{2} \nabla \log q_{T-t_k}(\mathbf{x}), \quad (\text{G.9})$$

$$\Psi_k(t, \mathbf{x}) = c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{x}) - \frac{1}{2} \nabla^2 \log q_{T-t_k}(\mathbf{x}). \quad (\text{G.10})$$

The interpolation operator can be expressed as:

$$F_t(\mathbf{x}) = \mathbf{x} + c_l(t - t_k) \mathbf{s}_\theta(T - t_k, \mathbf{x}), \quad (\text{G.11})$$

in this specific case. Recall that $g(t) = 1$ and $\mathbf{f}(T - t, \mathbf{x}) = 0$. Denote $F_t^{-1}(\mathbf{x})$ by \mathbf{z} . From (G.11) we have:

$$\frac{\partial}{\partial t} F(t, \mathbf{x}) = c_l \mathbf{s}_\theta(T - t_k, \mathbf{x}).$$

Thus we know that

$$\nabla \left[\frac{\partial}{\partial t} F \right] (t, F_t^{-1}(\mathbf{x})) = c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{z}). \quad (\text{G.12})$$

By Theorem 4.3, we know that:

$$\begin{aligned} \text{TV}(\mathbf{Y}_{T-\delta}, \hat{\mathbf{Y}}_{t_N}) &\leq \text{TV}(q_T, \pi_d) \\ &+ \sum_{k=0}^{N-1} \int_{t_k}^{t_{k+1}} \underbrace{\left[\sqrt{\mathbb{E}[\|\Phi_k(t, \mathbf{Y}_{t_k})\|^2]} \right]}_{J_1: \text{Score estimation error}} \underbrace{\sqrt{\int \left\| \frac{\nabla \log q(T-t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}}} \\ &+ \underbrace{\sqrt{\mathbb{E}[\text{tr}(\Psi_k(t, \mathbf{Y}_{t_k}))^2]} \sqrt{\int \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x}}}_{J_2: \text{Divergence estimation error}} \\ &+ \underbrace{\frac{1}{2} \int |(\nabla \log q_{T-t_k}(\mathbf{z}) - \nabla \log q_{T-t}(\mathbf{x})) \cdot \nabla \log q(T-t, \mathbf{x})| q(T-t, \mathbf{x}) d\mathbf{x}}_{J_3: \text{Score discretization error}} \\ &+ \underbrace{\frac{1}{2} \int |\text{tr}(\nabla^2 \log q(T-t_k, \mathbf{z}) - \nabla^2 \log q(T-t, \mathbf{x}))| q(T-t, \mathbf{x}) d\mathbf{x}}_{J_4: \text{Divergence discretization error}} \end{aligned}$$

$$\begin{aligned}
& + \underbrace{\max_{\mathbf{x}} \left| \text{tr} \left[c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{z}) \right] (\nabla F_t^{-1}(\mathbf{x}) - \mathbf{I}) \right|}_{J_5: \text{ Bias}} \text{d}t. \tag{G.13}
\end{aligned}$$

Since we assumed $\eta \leq \min\{1/(12L^2 T d^2), 1/(24L^2 R^2 d)\}$, we can easily verify that η is small enough to satisfy the condition of Lemma G.1 and Lemma G.2.

Bounding the score Estimation Error J_1 . For the first square root in J_1 , we know that:

$$\begin{aligned}
& \sqrt{\mathbb{E}[\|\Phi_k(t, \mathbf{Y}_{t_k})\|^2]} = \sqrt{\mathbb{E}\left[\|c_l \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \frac{1}{2} \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k})\|^2\right]} \\
& \leq \sqrt{\mathbb{E}\left[\left\|\frac{1}{2} \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \frac{1}{2} \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k})\right\| + (c_l - \frac{1}{2}) \|\mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\|\right]^2} \\
& \lesssim \sqrt{\mathbb{E}\left[\left\|\frac{1}{2} \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \frac{1}{2} \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k})\right\|^2\right]} + \sqrt{\mathbb{E}\left[(c_l - \frac{1}{2})^2 \|\mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\|^2\right]}. \tag{G.14}
\end{aligned}$$

Using Assumption 5.4, we have $\|\mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\| \leq L \|\mathbf{Y}_{t_k}\| + c$. Since $\mathbf{Y}_{t_k} = \mathbf{X}_0 + \sqrt{T - t_k} \mathbf{Z}$, we have:

$$\begin{aligned}
\mathbb{E}\|\mathbf{Y}_{t_k}\|^2 &= \mathbb{E}\|\mathbf{X}_0\|^2 + (T - t_k) \mathbb{E}\|\mathbf{Z}\|^2 \\
&\leq R^2 + (T - t_k)d.
\end{aligned}$$

Moreover, recall that we take $l = (T - t_k)/(t_{k+1} - t_k) \geq 1/\eta$. We can show that $0 \leq c_l - \frac{1}{2} \leq \eta$. Hence we have:

$$\begin{aligned}
\sqrt{\mathbb{E}\left[(c_l - \frac{1}{2})^2 \|\mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\|^2\right]} &\lesssim (c_l - \frac{1}{2}) \sqrt{\mathbb{E}\left[L^2 \|\mathbf{Y}_{t_k}\|^2 + c^2\right]} \\
&\leq \eta \sqrt{L^2 R^2 + L^2 (T - t_k)d + c^2}. \tag{G.15}
\end{aligned}$$

For the second term, since our η satisfies the condition of Lemma G.2 we can apply Lemma G.2 and obtain

$$\int_{\Omega_t} \left\| \frac{\nabla \log q(T - t, \mathbf{x}) p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) \text{d}\mathbf{x} \lesssim \frac{d}{T - t}. \tag{G.16}$$

Combining (G.15), (G.14) and (G.16), we have:

$$\begin{aligned}
J_1 &\lesssim \sqrt{\frac{d}{T - t}} \sqrt{\mathbb{E}\left[\|\mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k})\|^2\right]} \\
&\quad + \eta \sqrt{\frac{d}{T - t}} \sqrt{L^2 R^2 + L^2 (T - t_k)d + c^2}. \tag{G.17}
\end{aligned}$$

Bounding the Divergence Estimation Error J_2 . For the first term in J_2 , we know that:

$$\begin{aligned}
& \sqrt{\mathbb{E}\left[\text{tr} \left(c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \frac{1}{2} \nabla^2 \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right)^2\right]} \\
& \lesssim \sqrt{\mathbb{E}\left[\text{tr} \left(\frac{1}{2} \nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \frac{1}{2} \nabla^2 \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right)^2 + \left(c_l - \frac{1}{2}\right)^2 \left(\text{tr} \nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\right)^2\right]} \\
& \lesssim \sqrt{\mathbb{E}\left[\text{tr} \left(\nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right)^2\right]} + \underbrace{(c_l - \frac{1}{2}) \sqrt{\mathbb{E}\left[\left(\text{tr} \nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\right)^2\right]}}_{K_1}. \tag{G.18}
\end{aligned}$$

Using Assumption 5.4, we know that $\|\nabla \mathbf{s}_\theta(T - t_k, \mathbf{x})\|_2 \leq L$, thus $|\text{tr} \nabla \mathbf{s}_\theta(T - t_k, \mathbf{x})| \leq dL$. Same as we did with J_1 , we have $0 \leq c_l - \frac{1}{2} \leq \eta$. Moreover, K_1 can be further bounded by

$$\left(c_l - \frac{1}{2}\right) \sqrt{\mathbb{E}\left[\left(\text{tr} \nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k})\right)^2\right]} \leq \eta dL. \tag{G.19}$$

For the second term, since our η satisfies the condition of Lemma G.1, we can apply it and obtain

$$\int_{\Omega_t} \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim 1. \quad (\text{G.20})$$

Combining (G.18), (G.19) and (G.20) together, we have:

$$J_2 \lesssim \sqrt{\mathbb{E}_Q \operatorname{tr} \left(\nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q(T - t_k, \mathbf{Y}_{t_k}) \right)^2} + \eta dL. \quad (\text{G.21})$$

Bounding the Bias Term J_5 . We know that:

$$\begin{aligned} \nabla F_t^{-1}(\mathbf{x}) &= [(\nabla F_t)(F_t^{-1}(\mathbf{x}))]^{-1} \\ &= [\mathbf{I}_d + c_l(t - t_k) \nabla \mathbf{s}_\theta(T - t_k, F_t^{-1}(\mathbf{x}))]^{-1}. \end{aligned}$$

Using Assumption 5.4 and $t - t_k \leq \eta \leq 1$, we have:

$$\begin{aligned} \|(\nabla F_t)(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 &\leq (t - t_k) \|c_l \nabla \mathbf{s}_\theta(T - t_k, F_t^{-1}(\mathbf{x}))\|_2 \\ &\leq (t - t_k)(1 + L). \end{aligned}$$

Here we use $c_l \leq 1$. Let $\mathbf{A} = (\nabla F_t)(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d$. Then we have $\|\mathbf{A}\|_2 \leq (t - t_k)(1 + L)$. Thus we know that:

$$\begin{aligned} \|\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 &= \|(\mathbf{I}_d + \mathbf{A})^{-1} - \mathbf{I}_d\|_2 \\ &\stackrel{(i)}{=} \left\| \sum_{n=1}^{\infty} (-1)^n \mathbf{A}^n \right\|_2 \\ &\stackrel{(ii)}{\leq} \sum_{n=1}^{\infty} \|\mathbf{A}\|_2^n \\ &\stackrel{(iii)}{\leq} 2(t - t_k)(1 + L), \end{aligned}$$

here (i) holds because of the series expansion for $\|\mathbf{A}\|_2 < 1$. (ii) holds due to the Cauchy-Schwarz inequality. (iii) holds due to our assumption $\eta \leq 1/2(1 + L)$. Then we know that:

$$\begin{aligned} J_5 &= \max \left| \operatorname{tr} \left(c_l \nabla \mathbf{s}_\theta(T - t_k, \mathbf{z}) (\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d) \right) \right| \\ &\stackrel{(i)}{\leq} d \max \|\nabla \mathbf{s}_\theta(T - t_k, \mathbf{z}) (\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d)\|_2 \\ &\leq d \max \|\nabla \mathbf{s}_\theta(T - t_k, \mathbf{z})\|_2 \|\nabla(F_t^{-1}(\mathbf{x})) - \mathbf{I}_d\|_2 \\ &\leq dL(t - t_k)2(1 + L), \end{aligned}$$

here (i) holds due to $\operatorname{tr}(\mathbf{A}) \leq d\|\mathbf{A}\|_2$ and $c_l \leq 1$. Thus we have:

$$|J_5| \lesssim (t - t_k)dL^2. \quad (\text{G.22})$$

Estimating $\mathbf{z} - \mathbf{x}$. By Assumption 5.4, we have:

$$\begin{aligned} \|\mathbf{z} - \mathbf{x}\| &= \|F_t^{-1}(\mathbf{x}) - \mathbf{x}\| \\ &= (t - t_k) \|\mathbf{s}_\theta(T - t_k, F_t^{-1}(\mathbf{x}))\| \\ &\leq (t - t_k)L\|F_t^{-1}(\mathbf{x})\| + (t - t_k)c. \end{aligned}$$

Since we assume that $t - t_k \leq \eta \leq \min\{\frac{1}{4L}, \frac{1}{4c}\}$. From (G.7) we know that $\|F_t^{-1}(\mathbf{x})\| \leq 2\|\mathbf{x}\| + 1$. Hence we have:

$$\begin{aligned} \|\mathbf{z} - \mathbf{x}\| &\lesssim (t - t_k)L(2\|\mathbf{x}\| + 1) + (t - t_k)c \\ &\stackrel{(i)}{\lesssim} (t - t_k)L(\|\mathbf{x}\| + 1). \end{aligned} \quad (\text{G.23})$$

1728 **Bounding the Moments of \mathbf{Y}_t .**

1729 Before we start estimating J_3 and J_4 , we do some preparation work. Our goal is to estimate the
 1730 moments of \mathbf{Y}_t . Since we have $\mathbf{Y}_t = \mathbf{X}_{T-t} = \mathbf{X}_0 + \sqrt{T-t}\mathbf{Z}$, here $\mathbf{Z} \sim N(0, \mathbf{I}_d)$. We have
 1731

$$\begin{aligned} 1732 \mathbb{E}\|\mathbf{Y}_t\|^2 &= \mathbb{E}\|\mathbf{X}_0\|^2 + (T-t)\mathbb{E}\|\mathbf{Z}\|^2 \\ 1733 &\leq R^2 + (T-t)d. \end{aligned} \quad (\text{G.24})$$

1735 After computing the second order momentum, by the inequality $(\mathbb{E}\|\mathbf{Y}_t\|)^2 \leq \mathbb{E}\|\mathbf{Y}_t\|^2$, we know
 1736 that:

$$1737 \mathbb{E}\|\mathbf{Y}_t\| \lesssim R + \sqrt{T-t}\sqrt{d}. \quad (\text{G.25})$$

1738 Moreover, we consider the fourth order momentum as we will need to estimate it later: we have
 1739

$$\begin{aligned} 1740 \mathbb{E}\|\mathbf{Y}_t\|^4 &\lesssim \mathbb{E}\|\mathbf{X}_0\|^4 + (T-t)^2\mathbb{E}\|\mathbf{Z}\|^4 \\ 1741 &\lesssim R^4 + (T-t)^2d^2. \end{aligned} \quad (\text{G.26})$$

1743 After computing the forth order momentum, by the inequality $\mathbb{E}\|\mathbf{Y}_t\|^3 \leq \sqrt{\mathbb{E}\|\mathbf{Y}_t\|^2} \cdot \sqrt{\mathbb{E}\|\mathbf{Y}_t\|^4}$,
 1744 we know that:

$$1745 \mathbb{E}\|\mathbf{Y}_t\|^3 \lesssim R^3 + (T-t)^{3/2}d^{3/2}. \quad (\text{G.27})$$

1747 **Bounding Divergence Discretization Error J_4 .**

1748 Using Lemma G.3 and Lemma G.4, we know that

$$\begin{aligned} 1749 \left| \frac{\partial}{\partial t} \left(\text{tr}(\nabla^2 \log q(T-t, \mathbf{x})) \right) \right| &\lesssim \frac{d}{(T-t)^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{(T-t)^4}, \\ 1752 \|\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x}))\| &\leq \frac{R^3}{(T-t)^3}. \end{aligned}$$

1754 Thus we know that:

$$\begin{aligned} 1755 \left| \int_{\Omega_t} \text{tr}(\nabla^2 \log q(T-t_k, \mathbf{z}) - \nabla^2 \log q(T-t, \mathbf{x})) q(T-t, \mathbf{x}) d\mathbf{x} \right| \\ 1756 &\lesssim \int_{\Omega_t} \left((t-t_k) \left(\frac{d}{(T-t)^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{(T-t)^4} \right) + \frac{R^3}{(T-t)^3} \|\mathbf{x} - \mathbf{z}\| \right) q(T-t, \mathbf{x}) d\mathbf{x} \\ 1758 &\stackrel{(i)}{\lesssim} \int_{\Omega_t} \left((t-t_k) \left(\frac{d}{(T-t)^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{(T-t)^4} \right) + \frac{R^3}{(T-t)^3} (t-t_k) L(\|\mathbf{x}\| + 1) \right) q(T-t, \mathbf{x}) d\mathbf{x} \\ 1760 &\lesssim (t-t_k) \frac{d}{(T-t)^2} + (t-t_k) \frac{R^2}{(T-t)^4} \mathbb{E}(\|\mathbf{Y}_t\|^2 + R^2) + (t-t_k) \frac{R^3}{(T-t)^3} L \cdot \mathbb{E}(\|\mathbf{Y}_t\| + 1) \\ 1763 &\stackrel{(ii)}{\lesssim} (t-t_k) \frac{d}{(T-t)^2} + (t-t_k) \frac{R^4 + R^2(T-t)d}{(T-t)^4} + (t-t_k) \frac{LR^3}{(T-t)^3} (R + \sqrt{T-t}\sqrt{d} + 1). \end{aligned}$$

1767 Here (i) is due to (G.23) and (ii) is due to (G.24) and (G.25). Assuming $R \geq 1$, we can further
 1768 simplify its form:
 1769

$$1770 |J_4| \lesssim (t-t_k) \left[d \frac{1}{(T-t)^2} + d \frac{R^2}{(T-t)^3} + (1+L(T-t)) \frac{R^4}{(T-t)^4} + \sqrt{d} \frac{LR^3}{(T-t)^{5/2}} \right]. \quad (\text{G.28})$$

1772 **Bounding the Score Discretization Error J_3 .**

1773 Using Lemma G.4, we know that:
 1774

$$\begin{aligned} 1775 \|\nabla \log q(t, \mathbf{x})\| &\lesssim \frac{\|\mathbf{x}\| + R}{t}, \\ 1777 \|\nabla^2 \log q(t, \mathbf{x})\|_2 &\leq \frac{1}{t} + \frac{R^2}{t^2}. \end{aligned}$$

1779 Moreover, by Lemma G.3, we have:
 1780

$$1781 \left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| \lesssim \frac{\|\mathbf{x}\| + R}{t^2} + \frac{(\|\mathbf{x}\| + R)^3}{t^3}.$$

1782 Thus we know that:
 1783

$$\begin{aligned}
 & \left| \int_{\Omega_t} \left(\nabla \log q_{T-t_k}(\mathbf{z}) - \nabla \log q_{T-t}(\mathbf{x}) \right) \nabla \log q(T-t, \mathbf{x}) q(T-t, \mathbf{x}) d\mathbf{x} \right| \\
 & \lesssim \int_{\Omega_t} (t-t_k) \left(\frac{\|\mathbf{x}\| + R}{(T-t)^2} + \frac{(\|\mathbf{x}\| + R)^3}{(T-t)^3} \right) \frac{\|\mathbf{x}\| + R}{T-t} q(T-t, \mathbf{x}) d\mathbf{x} \\
 & + \int_{\Omega_t} \left(\frac{1}{T-t} + \frac{R^2}{(T-t)^2} \right) \|\mathbf{x} - \mathbf{z}\| \frac{\|\mathbf{x}\| + R}{T-t} q(T-t, \mathbf{x}) d\mathbf{x} \\
 & \lesssim (t-t_k) \frac{1}{(T-t)^3} \mathbb{E}(\|\mathbf{Y}_t\| + R)^2 + (t-t_k) \frac{1}{(T-t)^4} \mathbb{E}(\|\mathbf{Y}_t\| + R)^4 \\
 & + (t-t_k) \int_{\Omega_t} \left(\frac{1}{T-t} + \frac{R^2}{(T-t)^2} \right) L(\|\mathbf{x}\| + 1) \frac{\|\mathbf{x}\| + R}{T-t} q(T-t, \mathbf{x}) d\mathbf{x} \\
 & \stackrel{(i)}{\lesssim} (t-t_k) \frac{1}{(T-t)^3} (R^2 + (T-t)d) + (t-t_k) \frac{1}{(T-t)^4} (R^4 + (T-t)^2 d^2) \\
 & + (t-t_k) \frac{L}{(T-t)^2} (R^2 + (T-t)d) + (t-t_k) \frac{LR^2}{(T-t)^3} (R^2 + (T-t)d). \tag{G.29}
 \end{aligned}$$

1800 Here we use (G.26) and (G.24) to bound the second and fourth momentum of \mathbf{Y}_t , and (i) is due to:
 1801

$$\int_{\Omega_t} (\|\mathbf{x}\| + 1)(\|\mathbf{x}\| + R) q(T-t, \mathbf{x}) d\mathbf{x} \lesssim \mathbb{E}(\|\mathbf{Y}_t\| + R)^2 \lesssim R^2 + (T-t)d.$$

1802 Reorganizing terms in (G.29), since we assume that $L \geq 1$ and $R \geq 1$, we have:
 1803

$$\begin{aligned}
 |J_3| & \lesssim (t-t_k) \left[d^2 \frac{1}{(T-t)^2} + d \frac{1}{(T-t)^2} + d \frac{LR^2}{(T-t)^2} + d \frac{L}{T-t} \right. \\
 & \quad \left. + \frac{R^4}{(T-t)^4} + \frac{R^2}{(T-t)^3} + \frac{LR^2}{(T-t)^2} + \frac{LR^4}{(T-t)^3} \right] \\
 & \lesssim (t-t_k) \left[d^2 \frac{1}{(T-t)^2} + d \frac{LR^2}{(T-t)^2} + d \frac{L}{T-t} + \frac{R^4}{(T-t)^4} + \frac{LR^4}{(T-t)^3} + \frac{LR^2}{(T-t)^2} \right]. \tag{G.30}
 \end{aligned}$$

1814 Putting J_3 and J_4 Together.

1815 Combining (G.30) and (G.28) together, we have
 1816

$$\begin{aligned}
 |J_3| + |J_4| & \lesssim (t-t_k) \left[\underbrace{d^2 \frac{1}{(T-t)^2}}_{C_1} + \underbrace{d \frac{LR^2}{(T-t)^2}}_{C_2} + \underbrace{d \frac{L}{T-t}}_{C_3} + \underbrace{\frac{R^4}{(T-t)^4}}_{C_4} + \underbrace{\frac{LR^4}{(T-t)^3}}_{C_5} + \underbrace{\frac{LR^2}{(T-t)^2}}_{C_6} \right] \\
 & + (t-t_k) \left[\underbrace{d \frac{1}{(T-t)^2}}_{C_7} + \underbrace{d \frac{R^2}{(T-t)^3}}_{C_8} + \underbrace{(1+L(T-t)) \frac{R^4}{(T-t)^4}}_{C_9} + \underbrace{\sqrt{d} \frac{LR^3}{(T-t)^{5/2}}}_{C_{10}} \right] \\
 & \stackrel{(i)}{\lesssim} (t-t_k) [C_1 + C_2 + C_3 + C_4 + C_5 + C_8 + C_{10}] \\
 & = (t-t_k) \left[d^2 \frac{1}{(T-t)^2} + d \frac{LR^2}{(T-t)^2} + d \frac{L}{T-t} + \frac{R^4}{(T-t)^4} \right. \\
 & \quad \left. + \frac{LR^4}{(T-t)^3} + d \frac{R^2}{(T-t)^3} + \sqrt{d} \frac{LR^3}{(T-t)^{5/2}} \right], \tag{G.31}
 \end{aligned}$$

1831 where (i) is due to $C_7 \leq C_2$, $C_9 = C_4 + C_5$, $C_6 \leq C_2$.
 1832

Combining Everything Together.

1833 Plugging (G.17), (G.21), (G.22) and (G.31) into (G.13), and taking the integral, we can obtain the
 1834 following inequality
 1835

$$\text{TV}(\hat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \text{TV}(\pi_d, q_T)$$

$$\begin{aligned}
& + \underbrace{\sum_k (t_{k+1} - t_k) \sqrt{\mathbb{E}_Q \operatorname{tr} \left(\nabla \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla^2 \log q(T - t_k, \mathbf{Y}_{t_k}) \right)^2}}_{R_1} \\
& + \underbrace{\sum_k (t_{k+1} - t_k) \eta d L}_{R_2} \\
& + \underbrace{\sum_k (t_{k+1} - t_k) \sqrt{\frac{d}{T - t_{k+1}}} \sqrt{\mathbb{E}_Q \left\| \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2}}_{R_3} \\
& + \underbrace{\sum_k (t_{k+1} - t_k) \sqrt{\frac{d}{T - t_{k+1}}} \eta \sqrt{L^2 R^2 + L^2(T - t_k)d + c^2}}_{R_4} \\
& + \underbrace{\sum_k (t_{k+1} - t_k)^2 d L^2}_{R_5} \\
& + \sum_k (t_{k+1} - t_k)^2 \left[d^2 \frac{1}{(T - t)^2} + d \frac{LR^2}{(T - t)^2} + d \frac{L}{T - t} + \frac{R^4}{(T - t)^4} \right. \\
& \quad \left. + \frac{LR^4}{(T - t)^3} + d \frac{R^2}{(T - t)^3} + \sqrt{d} \frac{LR^3}{(T - t)^{5/2}} \right].
\end{aligned}$$

We further denote the last term by R_6 . We know that R_1 is ϵ_{div} . For R_3 , by Cauchy-Schwartz inequality, we have:

$$\begin{aligned}
R_3 & \leq \sqrt{\sum_k (t_{k+1} - t_k) \frac{d}{T - t_{k+1}}} \sqrt{\sum_k (t_{k+1} - t_k) \mathbb{E}_Q \left\| \mathbf{s}_\theta(T - t_k, \mathbf{Y}_{t_k}) - \nabla \log q_{T-t_k}(\mathbf{Y}_{t_k}) \right\|^2} \\
& \leq \sqrt{\eta N d} \cdot \epsilon_{\text{score}}.
\end{aligned}$$

Since we assume $L \geq 1$, for $R_2 + R_5$, we have

$$\sum_k (t_{k+1} - t_k) \eta d L + \sum_k (t_{k+1} - t_k)^2 d L^2 \lesssim \eta^2 N d L^2.$$

For the R_4 , we have

$$\begin{aligned}
& \sum_k (t_{k+1} - t_k) \sqrt{\frac{d}{T - t_{k+1}}} \eta \sqrt{L^2 R^2 + L^2(T - t_k)d + c^2} \\
& \stackrel{(i)}{\lesssim} \eta \sum_k (t_{k+1} - t_k) \left[\sqrt{\frac{d}{T - t_{k+1}}} \sqrt{L^2 R^2} + \sqrt{L^2(T - t_k)d} \right] \\
& \stackrel{(ii)}{\lesssim} \eta \sum_k \eta \sqrt{d} LR + (t - t_k) L d \\
& \leq \eta^2 N (\sqrt{d} LR + L d),
\end{aligned}$$

here we omit c in (i), and (ii) is because $\frac{T - t_k}{T - t_{k+1}} \leq 1 + \eta \leq 2$. Hence we know that

$$R_2 + R_4 + R_5 \lesssim \eta^2 N [L^2 d + LR \sqrt{d}].$$

Finally, for the R_6 , we know that:

$$R_6 \leq \eta^2 N \left[d^2 + LR^2 d + L d + \frac{R^4}{\delta^2} + \frac{LR^4}{\delta} + \frac{R^2}{\delta} d + LR^3 \sqrt{\frac{d}{\delta}} \right]$$

$$\lesssim \eta^2 N \left[d^2 + LR^2 d + \frac{R^2}{\delta} d + LR^3 \sqrt{\frac{d}{\delta}} + \frac{R^4}{\delta^2} + \frac{LR^4}{\delta} \right].$$

Putting together, we know that

$$\begin{aligned} \text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) &\lesssim \text{TV}(\pi_d, q_T) + \epsilon_{\text{div}} + \sqrt{\eta Nd} \cdot \epsilon_{\text{score}} \\ &\quad + \eta^2 N \left[d^2 + LR^2 d + L^2 d + \frac{R^2}{\delta} d + LR^3 \sqrt{\frac{d}{\delta}} + \frac{R^4}{\delta^2} + \frac{LR^4}{\delta} \right] \\ &\stackrel{(iii)}{\lesssim} \text{TV}(\pi_d, q_T) + \epsilon_{\text{div}} + \sqrt{\eta Nd} \cdot \epsilon_{\text{score}} + \eta^2 N \left[LR^4 d^2 + L^2 d + \frac{LR^4}{\delta^2} \right], \end{aligned}$$

here (iii) is due to $2LR^2 d \leq LR^4 d^2 + \frac{LR^4}{\delta^2}$, $2\frac{R^2}{\delta} d \leq LR^4 d^2 + \frac{LR^4}{\delta^2}$ and $LR^3 \sqrt{\frac{d}{\delta}} \leq LR^4 d^2 + \frac{LR^4}{\delta^2}$ (and $\delta < 1$ as well as $R \geq 1$ and $L \geq 1$). This completes the proof of Theorem 5.9. \square

G.2 PROOF OF LEMMA 5.10 AND COROLLARY 5.11

Proof of Lemma 5.10. Since $q_{t|0}(\mathbf{x}_t | \mathbf{x}_0) = \mathcal{N}(\mathbf{x}_t; \mathbf{x}_0, tI_d)$, we have

$$\text{KL}(q_{t|0}(\cdot | x_0) \| N(0, T\mathbf{I}_d)) = \frac{\|\mathbf{x}_0\|^2}{2T}. \quad (\text{G.32})$$

By the convexity of the KL divergence,

$$\begin{aligned} \text{KL}(q_T \| N(0, T\mathbf{I}_d)) &= \text{KL} \left(\int_{\mathbb{R}^d} q_{T|0}(\cdot | x_0) q_0(dx_0) \| N(0, T\mathbf{I}_d) \right) \\ &\leq \int_{\mathbb{R}^d} \text{KL}(q_{T|0}(\cdot | x_0) \| \pi_d) q_0(dx_0) \\ &= \frac{1}{2T} \mathbb{E}_{q_0} [\|\mathbf{X}_0\|^2] \\ &= \frac{d}{2T}. \end{aligned} \quad (\text{G.33})$$

where we have used that $\mathbb{E}_{q_0} [\|\mathbf{X}_0\|^2] = d$, since $\text{Cov}(q_0) = \mathbf{I}_d$. Applying Pinsker's inequality to (G.33), we have:

$$\text{TV}(\mathbf{X}_0 + N(0, T\mathbf{I}_d), N(0, T\mathbf{I}_d)) \leq \frac{\sqrt{d}}{\sqrt{2T}}. \quad (\text{G.34})$$

This completes the proof of Lemma 5.10. \square

Proof of Corollary 5.11. Please refer to Benton et al. (2024)'s Appendix D for a detailed derivation of the existence of such time schedule. Since we take $T = d/\epsilon^2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$, using $\epsilon \leq 1/L$, we can easily show that $\eta \leq \min\{1/(12L^2Td^2), 1/(24L^2R^2d)\}$. Hence by Theorem 5.5 and Lemma 5.6, we have:

$$\text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \frac{\sqrt{d}}{\sqrt{T}} + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N \left[LR^4 d^2 + L^2 d + LR^4 \frac{1}{\delta^2} \right].$$

Since we take $\delta = 1/d$ and assume that $d \geq L/R^4$, we can further simply the bound of TV to:

$$\text{TV}(\widehat{\mathbf{Y}}_{T-\delta}, \mathbf{Y}_{T-\delta}) \lesssim \frac{\sqrt{d}}{\sqrt{T}} + \epsilon_{\text{div}} + \sqrt{d} \sqrt{\eta N} \epsilon_{\text{score}} + \eta^2 N L R^4 d^2.$$

Since $T = d/\epsilon^2$ and $N = LR^4 \Theta(d^2(T + \log(1/\delta))^2/\epsilon)$, we know that

$$\begin{aligned} \frac{\sqrt{d}}{\sqrt{T}} &\leq \epsilon, \\ \eta^2 N L R^4 d^2 &= \frac{1}{N} (T + \log(1/\delta))^2 L R^4 d^2 \leq \epsilon. \end{aligned}$$

Recall that we assume sufficiently small score estimation and divergence error, this completes the proof of Corollary 5.11. \square

1944 H LIPSCHITZNESS OF SCORE FUNCTIONS

1945 In this section, we consider the forward process by gradually adding noise to data distribution, i.e.,

$$1947 \quad \mathbf{X}_t = f(t)\mathbf{X}_{\text{data}} + g(t)\mathbf{Z}, \quad (\text{H.1})$$

1948 where $\mathbf{Z} \sim \mathcal{N}(0, I)$ is a standard normal random variable independent of \mathbf{X}_{data} , and $f(t)$ and $g(t)$
 1949 are scaling functions determined by the forward process that modulate the influence of the data
 1950 and noise over time. Both VE and VP forward processes can be represented in this form. The
 1951 distribution of \mathbf{X}_t is denoted by $q(t, \mathbf{x})$. We will show the Lipschitzness property of $\nabla \log q(t, \mathbf{x})$
 1952 under Assumption 5.1.

1953 H.1 GENERAL RESULTS

1954 The following lemma regards the upper bound of the derivatives concerning \mathbf{x} .

1955 **Lemma H.1.** Suppose Assumption 5.1 holds. Let \mathbf{X}_t be the forward process defined in (H.1), and
 1956 its distribution is denoted by $q(t, \mathbf{x})$. We have the following inequalities:

$$\begin{aligned} 1958 \quad & \|\nabla \log q(t, \mathbf{x})\| \leq \frac{\|\mathbf{x}\| + f(t)R}{g(t)^2}, \\ 1959 \quad & \|\nabla^2 \log q(t, \mathbf{x})\|_2 \leq \frac{1}{g(t)^2} \left(1 + \frac{f(t)^2}{g(t)^2} 2R^2\right), \\ 1960 \quad & \|\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x}))\| \leq \frac{6f(t)^3}{g(t)^6} R^3. \end{aligned}$$

1965 *Proof of Lemma H.1.* We prove the lemma by gradually computing the higher-order derivatives of
 1966 the score function step by step. A key point in our proof is that we bring the terms involving \mathbf{x}
 1967 outside of the integral and canceling them, resulting in an expression where \mathbf{x} only appears in the
 1968 function's evaluation. Firstly, we derive the form of our score function $\nabla \ln q(t, \mathbf{x})$.

1969 Expression of the Score Function.

1970 To start with, using the expression of the forward process \mathbf{X}_t in (H.1), we represent $q(t, \mathbf{x})$ as the
 1971 convolution of the data distribution and Gaussian distribution. Specifically, we have:

$$\begin{aligned} 1972 \quad q(t, \mathbf{x}) &= \int_{\mathbb{R}^d} p_{f(t)} \mathbf{x}_0(\mathbf{y}) p_{g(t)} \mathbf{Z}(\mathbf{x} - \mathbf{y}) d\mathbf{y} \\ 1973 \quad &= \int_{\mathbb{R}^d} \frac{1}{f(t)} p_{\mathbf{x}_0} \left(\frac{\mathbf{y}}{f(t)} \right) p_{g(t)} \mathbf{Z}(\mathbf{x} - \mathbf{y}) d\mathbf{y} \\ 1974 \quad &= \int_{\mathbb{R}^d} p_{\mathbf{x}_0}(\mathbf{y}) p_{g(t)} \mathbf{Z}(\mathbf{x} - f(t)\mathbf{y}) d\mathbf{y} \\ 1975 \quad &= \int_{\mathbb{R}^d} q_0(\mathbf{y}) p_{\mathbf{Z}} \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) \frac{1}{g(t)} d\mathbf{y} \\ 1976 \quad &= \int_{\mathbb{R}^d} q_0(\mathbf{y}) \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) \frac{1}{g(t)} d\mathbf{y}, \end{aligned} \quad (\text{H.2})$$

1977 where we represent the probability density function of standard Gaussian distribution as $\phi(\mathbf{x}) = (e^{-\|\mathbf{x}\|^2/2})/(\sqrt{2\pi})^d$. Besides, in the second and fourth equations, we use the change of variable
 1978 formula of the density function, i.e., $p_{a\mathbf{X}}(\mathbf{x}) = p_{\mathbf{X}}(\mathbf{x}/a)/a$, where \mathbf{X} is a random variable and a is
 1979 a constant. By directly taking the gradient, we have the following property of $\phi(\cdot)$:

$$1980 \quad \nabla \phi(\mathbf{x}) = -\phi(\mathbf{x})\mathbf{x}.$$

1981 Moreover, by the chain rule, we can calculate the gradient with respect to \mathbf{x} :

$$\begin{aligned} 1982 \quad \nabla_{\mathbf{x}} \left[\phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) \right] &= -\frac{1}{g(t)} \left[\phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) \frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right] \\ 1983 \quad &= -\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2} \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right). \end{aligned} \quad (\text{H.3})$$

1984 Thus, we can express the score function as follows:

$$1985 \quad \nabla \log q(t, \mathbf{x}) = \frac{\nabla q(t, \mathbf{x})}{q(t, \mathbf{x})}$$

$$\begin{aligned}
& \stackrel{(i)}{=} \frac{\nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) \frac{1}{g(t)} d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) \frac{1}{g(t)} d\mathbf{y}} \\
& = \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \nabla_{\mathbf{x}} \left[\phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) \right] d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \\
& \stackrel{(ii)}{=} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(-\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)^2} \right) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \\
& = -\frac{1}{g(t)^2} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}, \tag{H.4}
\end{aligned}$$

where (i) holds due to (H.2), and (ii) holds due to (H.3).

Under Assumption 5.1, we have $p_{\text{data}}(\mathbf{y}) = 0$ for $\|\mathbf{y}\| > R$. Thus, we have

$$\begin{aligned}
\|\nabla \log q(t, \mathbf{x})\| &= \frac{1}{g(t)^2} \frac{\left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y} \right\|}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \\
&\leq \frac{1}{g(t)^2} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{x} - f(t)\mathbf{y}\| \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \\
&\leq \frac{1}{g(t)^2} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\|\mathbf{x}\| + f(t)\|\mathbf{y}\|) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \\
&\leq \frac{\|\mathbf{x}\| + f(t)R}{g(t)^2}.
\end{aligned}$$

This shows that, under Assumption 5.1, the norm of the score function can be bounded by a linear function.

First-order Derivative of the Score Function.

We proceed by computing the Jacobian matrix of $\nabla \log q(t, \mathbf{x})$. Directly calculating the gradient of (H.6), we have

$$\begin{aligned}
\nabla^2 \log q(t, \mathbf{x}) &= \nabla \left[-\frac{1}{g(t)^2} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}} \right] \\
&= -\frac{1}{g(t)^2} \frac{1}{\Phi(\mathbf{x})^2} \left[\underbrace{\nabla \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y} \right)}_{I_1} \Phi(\mathbf{x}) \right. \\
&\quad \left. - \underbrace{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y} \left(\nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y} \right)^\top}_{I_2} \right], \tag{H.5}
\end{aligned}$$

where we use the shorthand expression $\Phi(\mathbf{x}) = \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}$, and the last equation holds due to the gradient formula of vector-valued functions, i.e., $\nabla_{\mathbf{g}(\mathbf{x})} \frac{\nabla f(\mathbf{x})}{g(\mathbf{x})} = \frac{\nabla f(\mathbf{x})}{g(\mathbf{x})} - \frac{f(\mathbf{x}) \nabla g(\mathbf{x})^\top}{g(\mathbf{x})^2}$.

Firstly, we compute I_1 :

$$\begin{aligned}
I_1 &= \nabla \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y} \right) \\
&= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(\phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) \mathbf{I} + (\mathbf{x} - f(t)\mathbf{y}) \nabla_{\mathbf{x}} \left[\phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) \right]^\top \right) d\mathbf{y} \\
&\stackrel{(i)}{=} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(\phi \mathbf{I} + (\mathbf{x} - f(t)\mathbf{y}) \left(-\phi \frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2} \right)^\top \right) d\mathbf{y}
\end{aligned}$$

$$\begin{aligned}
&= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(\mathbf{I} - \frac{(\mathbf{x} - f(t)\mathbf{y})(\mathbf{x} - f(t)\mathbf{y})^\top}{g(t)^2} \right) \phi d\mathbf{y} \\
&= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \mathbf{I} - \underbrace{\frac{\mathbf{x}\mathbf{x}^\top}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}}_{I_{1,1}} + \underbrace{\frac{f(t)}{g(t)^2} \mathbf{x} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y}^\top \phi d\mathbf{y}}_{I_{1,2}} \\
&\quad + \underbrace{\frac{f(t)}{g(t)^2} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \mathbf{x}^\top}_{I_{1,3}} - \frac{f(t)^2}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y}, \tag{H.6}
\end{aligned}$$

where we use the shorthand expression $\phi = \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right)$ and (i) holds due to (H.3).

Next, we compute I_2 :

$$\begin{aligned}
I_2 &= \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) d\mathbf{y} \right] \left[\nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) d\mathbf{y} \right]^\top \\
&= \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) d\mathbf{y} \right] \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(-\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2} \right) \phi \left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)} \right) d\mathbf{y} \right]^\top \\
&= -\frac{1}{g(t)^2} \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi d\mathbf{y} \right] \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y})^\top \phi d\mathbf{y} \right] \\
&= -\underbrace{\frac{\mathbf{x}\mathbf{x}^\top}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}}_{I_{2,1}} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} + \underbrace{\frac{f(t)}{g(t)^2} \mathbf{x} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y}^\top \phi d\mathbf{y}}_{I_{2,2}} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \\
&\quad + \underbrace{\frac{f(t)}{g(t)^2} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \mathbf{x}^\top}_{I_{2,3}} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} - \frac{f(t)^2}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y}^\top \phi d\mathbf{y}, \tag{H.7}
\end{aligned}$$

where the second inequality holds due to (H.3), the third inequality holds due to the shorthand expression $\phi = \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right)$. Substituting (H.6) and (H.7) into (H.5), notice that $\Phi(\mathbf{x}) = \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi \left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)} \right) d\mathbf{y}$, $I_{1,1} = I_{2,1}$, $I_{1,2} = I_{2,2}$ and $I_{1,3} = I_{2,3}$. Thus, we have the following equation:

$$\begin{aligned}
\nabla^2 \log q(t, \mathbf{x}) &= -\frac{1}{g(t)^2} \left[\mathbf{I} - \frac{f(t)^2}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \right. \right. \\
&\quad \left. \left. - \frac{\left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right] \left[\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y}^\top \phi d\mathbf{y} \right]}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right) \right]. \tag{H.8}
\end{aligned}$$

Using Assumption 5.1, when $p_{\text{data}}(\mathbf{y}) \neq 0$, we have $\|\mathbf{y} \mathbf{y}^\top\|_2 = \|\mathbf{y}\|^2 \leq R^2$. Moreover, we have

$$\begin{aligned}
\left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y}^\top \phi d\mathbf{y} \right\|_2 &= \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2 \\
&\leq \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \right)^2 \\
&\leq R^2 \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2.
\end{aligned}$$

Substituting into (H.8), we have

$$\|\nabla^2 \log q(t, \mathbf{x})\|_2 \leq \frac{1}{g(t)^2} \left(1 + \frac{f(t)^2}{g(t)^2} 2R^2 \right).$$

This implies that $\nabla \log q(t, \mathbf{x})$ is Lipschitz.

Derivatives of the Divergence of the Score Function.

Next, we consider the divergence of the score function, i.e.,

$$\nabla \cdot [\nabla \log q(t, \mathbf{x})] = \text{tr}(\nabla^2 \log q(t, \mathbf{x})).$$

To start with, using (H.8), we have:

$$\text{tr}(\nabla^2 \log q(t, \mathbf{x})) = -\frac{d}{g(t)^2} + \frac{f(t)^2}{g(t)^4} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2} \right), \quad (\text{H.9})$$

where we use the fact $\text{tr}(\mathbf{x}\mathbf{x}^\top) = \|\mathbf{x}\|^2$. Directly computing the gradient, we have

$$\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x})) = \frac{f(t)^2}{g(t)^4} \left[\underbrace{\nabla \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}}}_{J_1} - \underbrace{\nabla \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2}}_{J_2} \right].$$

Firstly, for J_1 , we have

$$J_1 = \frac{\nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} - \nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2}. \quad (\text{H.10})$$

We calculate the two gradients separately. First, we have

$$\begin{aligned} \nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} &= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \nabla \phi d\mathbf{y} \\ &= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \left(-\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2}\right) \phi d\mathbf{y} \\ &= -\frac{\mathbf{x}}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} + \frac{f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \mathbf{y} \phi d\mathbf{y}, \end{aligned} \quad (\text{H.11})$$

where the second equality holds due to (H.3). Next, we have

$$\begin{aligned} \nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} &= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \nabla \phi d\mathbf{y} \\ &= \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left(-\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2}\right) \phi d\mathbf{y} \\ &= -\frac{\mathbf{x}}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} + \frac{f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}. \end{aligned} \quad (\text{H.12})$$

Substituting (H.11) and (H.12) into (H.10), we have

$$\begin{aligned} J_1 &= \frac{\left(-\frac{\mathbf{x}}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} + \frac{f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \mathbf{y} \phi d\mathbf{y}\right) \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2} \\ &\quad - \frac{\left(-\frac{\mathbf{x}}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} + \frac{f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\right) \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2} \\ &= \frac{f(t)}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \mathbf{y} \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2} \right). \end{aligned}$$

Next, for J_2 , we have:

$$J_2 = \frac{\nabla \|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2 \cdot \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2 - \nabla \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^2 \cdot \|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}\right)^4}. \quad (\text{H.13})$$

Again, we calculate the two gradients as follows. First, we have:

$$\nabla \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2 = \nabla \sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \right)^2$$

$$\begin{aligned}
&= \sum_{i=1}^d 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \nabla \phi d\mathbf{y} \\
&\stackrel{(i)}{=} \sum_{i=1}^d 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \left(-\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)^2} \right) \phi d\mathbf{y} \\
&= -\frac{2\mathbf{x}}{g(t)^2} \sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \right)^2 + \frac{2f(t)}{g(t)^2} \sum_{i=1}^d \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \mathbf{y} \phi d\mathbf{y} \\
&\stackrel{(ii)}{=} -\frac{2\mathbf{x}}{g(t)^2} \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2 + \frac{2f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}, \quad (\text{H.14})
\end{aligned}$$

where (i) holds due to (H.3), and (ii) holds due to the coordinate expression of matrix multiplication. Moreover, we have

$$\begin{aligned}
\nabla \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2 &= 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \nabla \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \\
&\stackrel{(i)}{=} 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \left(-\frac{\mathbf{x}}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} + \frac{f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \\
&= -\frac{2\mathbf{x}}{g(t)^2} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2 + \frac{2f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}, \quad (\text{H.15})
\end{aligned}$$

where (i) holds due to (H.12). Substituting (H.14) and (H.15) into (H.13), we have

$$\begin{aligned}
J_2 &= \frac{\left(-\frac{2\mathbf{x}}{g(t)^2} \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2 + \frac{2f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^4} \\
&\quad - \frac{\left(-\frac{2\mathbf{x}}{g(t)^2} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2 + \frac{2f(t)}{g(t)^2} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \cdot \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^4} \\
&= \frac{2f(t)}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} + \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \cdot \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^3} \right).
\end{aligned}$$

We can conclude with

$$\begin{aligned}
\|J_1\| &\leq \frac{f(t)}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \|\mathbf{y}\| \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right) \\
&\leq \frac{f(t)}{g(t)^2} \left(R^3 + R \cdot R^2 \right) \\
&= \frac{2f(t)}{g(t)^2} R^3.
\end{aligned}$$

Moreover, we have

$$\begin{aligned}
\|J_2\| &\leq \frac{2f(t)}{g(t)^2} \left(\frac{\left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \mathbf{y}^\top \phi d\mathbf{y} \right\|_2 \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right. \\
&\quad \left. + \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \cdot \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^3} \right) \\
&\leq \frac{2f(t)}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y} \mathbf{y}^\top\|_2 \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right. \\
&\quad \left. + \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \cdot \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \right)^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^3} \right)
\end{aligned}$$

$$\begin{aligned}
&\stackrel{(i)}{=} \frac{2f(t)}{g(t)^2} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right. \\
&\quad \left. + \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \cdot \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\| \phi d\mathbf{y} \right)^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^3} \right) \\
&\stackrel{(ii)}{\leq} \frac{2f(t)}{g(t)^2} \left(R^2 \cdot R + R \cdot R^2 \right) = \frac{4f(t)}{g(t)^2} R^3,
\end{aligned}$$

where (i) holds due to the fact $\|\mathbf{y}\mathbf{y}^\top\|_2 = \|\mathbf{y}\|^2$ for any $\mathbf{y} \in \mathbb{R}^d$, and (ii) holds due to Assumption 5.1. Putting everything together, we know that

$$\|\nabla \text{tr}(\nabla^2 \log q(t, \mathbf{x}))\| \leq \frac{6f(t)^3}{g(t)^6} R^3.$$

This completes the proof of Lemma H.1. \square

The following lemma considers the upper bounds for the time-derivative of our score function. We use similar techniques to as we obtain in the proof of Lemma H.1.

Lemma H.2. Suppose Assumption 5.1 holds. Let X_t be the forward process defined in (H.1), and its distribution is denoted by $q(t, \mathbf{x})$. Then we have:

$$\begin{aligned}
\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| &\leq \frac{2g'(t)}{g(t)^3} \left(\|\mathbf{x}\| + |f(t)|R \right) + \frac{|f'(t)|R}{g(t)^2} \\
&\quad + \frac{2g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2)^2 + 2|g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^3}{g(t)^5}, \\
\left| \frac{\partial}{\partial t} (\text{tr}(\nabla^2 \log q(t, \mathbf{x}))) \right| &\leq \frac{2d \cdot |g'(t)|}{g(t)^3} + \frac{4f(t) \cdot |f'(t)g(t) - 2f(t)g'(t)|}{g(t)^5} R^2 \\
&\quad + 6R^2 \frac{f(t)^2}{g(t)^4} \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \right).
\end{aligned}$$

Proof of Lemma H.2. To start with, we first compute the time-derivative of $\phi(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)})$. Using the chain rule, we have

$$\begin{aligned}
\frac{\partial}{\partial t} \phi\left(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\right) &= \frac{\partial}{\partial t} \left(\frac{1}{(\sqrt{2\pi})^d} e^{-\frac{\|\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\|^2}{2}} \right) \\
&= \frac{1}{(\sqrt{2\pi})^d} e^{-\frac{\|\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)}\|^2}{2}} \left(-\frac{1}{2} \right) \frac{\partial}{\partial t} \left\| \frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)} \right\|^2. \quad (\text{H.16})
\end{aligned}$$

Since we know that:

$$\begin{aligned}
\frac{\partial}{\partial t} \left\| \frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)} \right\|^2 &= \frac{\partial}{\partial t} \frac{\|\mathbf{x}\|^2 - 2f(t)\mathbf{x}^\top\mathbf{y} + f(t)^2\|\mathbf{y}\|^2}{g(t)^2} \\
&= \frac{(-2f'(t)\mathbf{x}^\top\mathbf{y} + 2f(t)f'(\mathbf{y})\|\mathbf{y}\|^2)g(t)^2 - 2\|\mathbf{x}-f(t)\mathbf{y}\|^2g(t)g'(t)}{g(t)^4} \\
&= -\frac{2}{g(t)^3} \left(g(t)f'(t)(\mathbf{x}^\top\mathbf{y} - f(t)\|\mathbf{y}\|^2) + g'(t)\|\mathbf{x}-f(t)\mathbf{y}\|^2 \right). \quad (\text{H.17})
\end{aligned}$$

Substituting (H.17) into (H.16), and recall our shorthand expression $\phi = \phi(\frac{\mathbf{x}-f(t)\mathbf{y}}{g(t)})$, we know that:

$$\frac{\partial}{\partial t} \phi = \frac{g(t)f'(t)(\mathbf{x}^\top\mathbf{y} - f(t)\|\mathbf{y}\|^2) + g'(t)\|\mathbf{x}-f(t)\mathbf{y}\|^2}{g(t)^3} \phi. \quad (\text{H.18})$$

Therefore, assuming $\|\mathbf{y}\| \leq R$, we have

$$\left| \frac{\partial}{\partial t} \phi \right| = \left| \frac{g(t)f'(t)(\mathbf{x}^\top\mathbf{y} - f(t)\|\mathbf{y}\|^2) + g'(t)\|\mathbf{x}-f(t)\mathbf{y}\|^2}{g(t)^3} \right| \phi$$

$$\begin{aligned}
&\leq \frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\| \|\mathbf{y}\| + f(t) \|\mathbf{y}\|^2) + |g'(t)| \cdot \|\mathbf{x} - f(t)\mathbf{y}\|^2}{g(t)^3} \phi \\
&\leq \frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\| R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \phi. \tag{H.19}
\end{aligned}$$

Time-Derivative of the Score Function.

By (H.4), we know:

$$\nabla \log q(t, \mathbf{x}) = -\frac{1}{g(t)^2} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi\left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi\left(\frac{\mathbf{x} - f(t)\mathbf{y}}{g(t)}\right) d\mathbf{y}}.$$

Taking the derivative with respect to t , we have:

$$\begin{aligned}
\frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) &= \frac{2g'(t)}{g(t)^3} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \\
&\quad - \frac{1}{g(t)^2} \underbrace{\left(\frac{\partial}{\partial t} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \right)}_{L_1}.
\end{aligned}$$

Using Assumption 5.1, we can easily see that:

$$\left\| \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \right\| \leq \|\mathbf{x}\| + |f(t)|R. \tag{H.20}$$

For L_1 , we have:

$$L_1 = \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) ((\mathbf{x} - f(t)\mathbf{y}) \phi_t - f'(t)\mathbf{y}\phi) d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)} - \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) (\mathbf{x} - f(t)\mathbf{y}) \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_t d\mathbf{y}}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2}.$$

From (H.19) and Assumption 5.1, we know that:

$$\|L_1\| \leq \frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} 2(\|\mathbf{x}\| + |f(t)|R) + |f'(t)|R. \tag{H.21}$$

Combining (H.20) and (H.21), we have:

$$\begin{aligned}
\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| &= \frac{2g'(t)}{g(t)^3} (\|\mathbf{x}\| + |f(t)|R) \\
&\quad + \frac{1}{g(t)^2} \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} 2(\|\mathbf{x}\| + |f(t)|R) + |f'(t)|R \right) \\
&= \frac{2g'(t)}{g(t)^3} (\|\mathbf{x}\| + |f(t)|R) + \frac{2g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2)^2 + 2|g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^3}{g(t)^5} \\
&\quad + \frac{|f'(t)|R}{g(t)^2}.
\end{aligned}$$

Time-Derivative of the Divergence of the Score Function.

In this section, we consider the t -derivative of the divergence of the score function, i.e.,

$$\frac{\partial}{\partial t} (\nabla \cdot [\nabla \log q(t, \mathbf{x})]) = \frac{\partial}{\partial t} (\text{tr} (\nabla^2 \log q(t, \mathbf{x}))).$$

To start with, using (H.9), we have:

$$\text{tr} (\nabla^2 \log q(t, \mathbf{x})) = -\frac{d}{g(t)^2} + \frac{f(t)^2}{g(t)^4} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right). \tag{H.22}$$

Taking the derivative with respect to t , we have:

$$\begin{aligned} \frac{\partial}{\partial t} \left(\text{tr} (\nabla^2 \log q(t, \mathbf{x})) \right) &= \frac{2d \cdot g'(t)}{g(t)^3} \\ &+ \underbrace{\frac{2f(t)f'(t)g(t) - 4f(t)^2g'(t)}{g(t)^5} \left(\frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2} \right)}_{K_1} \\ &+ \underbrace{\frac{f(t)^2}{g(t)^4} \left(\frac{\partial}{\partial t} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\partial}{\partial t} \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2} \right)}_{K_2} - \underbrace{\frac{\partial}{\partial t} \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2}}_{K_3}. \end{aligned} \quad (\text{H.23})$$

For K_1 , using Assumption 5.1, we have

$$|K_1| \leq \left| \frac{2f(t)f'(t)g(t) - 4f(t)^2g'(t)}{g(t)^5} (R^2 + R^2) \right| = \frac{4f(t) \cdot |f'(t)g(t) - 2f(t)g'(t)|}{g(t)^5} R^2. \quad (\text{H.24})$$

For K_2 , Using the definition of K_2 in (H.23), we have

$$\begin{aligned} |K_2| &= \left| \frac{\partial}{\partial t} \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \right| \\ &= \left| \frac{\frac{\partial}{\partial t} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} - \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} \cdot \left(\frac{\partial}{\partial t} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right| \\ &\leq \left| \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \frac{\partial}{\partial t} \phi d\mathbf{y}}{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y}} \right| + \left| \frac{\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} \cdot \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \frac{\partial}{\partial t} \phi d\mathbf{y} \right)}{\left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2} \right| \\ &\leq 2R^2 \frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3}, \end{aligned} \quad (\text{H.25})$$

where the last inequality holds due to (H.19) and Assumption 5.1.

For K_3 , we have

$$\begin{aligned} |K_3| &= \left| \frac{\partial}{\partial t} \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2} \right| \\ &\leq \left| \frac{\frac{\partial}{\partial t} \|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2} - \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2 \cdot \frac{\partial}{\partial t} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^4} \right| \\ &\leq \left| \frac{\frac{\partial}{\partial t} \|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^2} \right| + \left| \frac{\|\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y}\|^2 \cdot \frac{\partial}{\partial t} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2}{(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y})^4} \right|. \end{aligned} \quad (\text{H.26})$$

Moreover, we have:

$$\begin{aligned} \left| \frac{\partial}{\partial t} \left\| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right\|^2 \right| &= \left| \frac{\partial}{\partial t} \sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \right)^2 \right| \\ &= \left| \sum_{i=1}^d 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \phi d\mathbf{y} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) y_i \frac{\partial}{\partial t} \phi d\mathbf{y} \right| \\ &\leq \sum_{i=1}^d 2 \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) |y_i| \phi d\mathbf{y} \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) |y_i| \left| \frac{\partial}{\partial t} \phi \right| d\mathbf{y} \\ &\leq 2 \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \right) \sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) |y_i| \phi d\mathbf{y} \right)^2. \end{aligned}$$

Using the Cauchy-Schwarz inequality, we have

$$\sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) |y_i| \phi d\mathbf{y} \right)^2 \leq \sum_{i=1}^d \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right) \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) |y_i|^2 \phi d\mathbf{y} \right)$$

$$\begin{aligned}
&= \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right) \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \|\mathbf{y}\|^2 \phi d\mathbf{y} \right) \\
&\leq R^2 \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2,
\end{aligned}$$

where the last inequality holds due to Assumption 5.1. Therefore, we have

$$\begin{aligned}
&\left| \frac{\partial}{\partial t} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \mathbf{y} \phi d\mathbf{y} \right) \right|^2 \\
&\leq 2R^2 \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \right) \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2. \tag{H.27}
\end{aligned}$$

Moreover, we have:

$$\begin{aligned}
&\left| \frac{\partial}{\partial t} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2 \right| = 2 \left| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \cdot \frac{\partial}{\partial t} \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right| \\
&\leq 2 \left| \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right| \cdot \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \left| \frac{\partial}{\partial t} \phi \right| d\mathbf{y} \\
&\leq 2 \frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi d\mathbf{y} \right)^2, \tag{H.28}
\end{aligned}$$

where the last inequality holds due to (H.18). Substituting (H.27) and (H.28) into (H.26), and using Assumption 5.1, we have:

$$|K_3| \leq 4R^2 \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \right). \tag{H.29}$$

Combining (H.23), (H.24), (H.25) and (H.29), we have the following inequality:

$$\begin{aligned}
&\left| \frac{\partial}{\partial t} \left(\text{tr}(\nabla^2 \log q(t, \mathbf{x})) \right) \right| \leq \frac{2d \cdot |g'(t)|}{g(t)^3} + \frac{4f(t) \cdot |f'(t)g(t) - 2f(t)g'(t)|}{g(t)^5} R^2 \\
&\quad + 6R^2 \frac{f(t)^2}{g(t)^4} \left(\frac{g(t)|f'(t)| \cdot (\|\mathbf{x}\|R + f(t)R^2) + |g'(t)| \cdot (\|\mathbf{x}\| + f(t)R)^2}{g(t)^3} \right). \tag{H.30}
\end{aligned}$$

This completes the proof of Lemma H.2. \square

H.2 PROOF OF LEMMAS F.3 AND F.4

Proof of Lemma F.3. Since $\mathbf{X}_t = e^{-t} \mathbf{X}_0 + N(0, (1 - e^{-2t}) \mathbf{I}_d)$, we have $f(t) = e^{-t}$ and $g(t) = \sqrt{1 - e^{-2t}}$ in this case. When $t < 1$, we know that $f(t) = \Theta(1)$, $f'(t) = \Theta(1)$, $g(t) = \Theta(\sqrt{t})$, $g'(t) = \Theta(\frac{1}{\sqrt{t}})$. Hence by Lemma H.2, we have:

$$\begin{aligned}
&\left| \frac{\partial}{\partial t} \left(\text{tr}(\nabla^2 \log q(t, \mathbf{x})) \right) \right| \lesssim \frac{d}{t^2} + \frac{R^2}{t^3} + R^2 \frac{1}{t^2} \frac{\sqrt{t}(\|\mathbf{x}\| + R)R + \frac{1}{\sqrt{t}}(\|\mathbf{x}\| + R)^2}{\sqrt{t}^3} \\
&\quad = \frac{d}{t^2} + \frac{R^2}{t^3} + \frac{R^3(\|\mathbf{x}\| + R)}{t^3} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4} \\
&\quad \stackrel{(i)}{\lesssim} \frac{d}{t^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4}, \quad t < 1.
\end{aligned}$$

Here in (i) we use $2 \frac{R^2}{t^3} \leq \frac{1}{t^2} + \frac{R^4}{t^4} \leq \frac{d}{t^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4}$. We next consider the case $t \geq 1$:

Denote e^{-t} by a , then we know that $f(t) = \Theta(a)$, $f'(t) = \Theta(a)$, $g(t) = \Theta(1)$, $g'(t) = \Theta(a^2)$. By, Lemma H.2, we have:

$$\left| \frac{\partial}{\partial t} \left(\text{tr}(\nabla^2 \log q(t, \mathbf{x})) \right) \right| \lesssim a^2 d + R^2 a^2 + R^2 a^2 \left(a(\|\mathbf{x}\| + aR)R + a^2(\|\mathbf{x}\| + aR)^2 \right)$$

$$\begin{aligned}
&\stackrel{(ii)}{\leq} a^2d + R^2a^2 + R^2a^2((\|\mathbf{x}\| + R)R + (\|\mathbf{x}\| + R)^2) \\
&\lesssim a^2d + a^2R^2 + a^2R^2(\|\mathbf{x}\| + R)^2 \\
&\stackrel{(iii)}{\lesssim} a^2d + a^2R^2(\|\mathbf{x}\| + R)^2 \\
&\leq d + R^2(\|\mathbf{x}\| + R)^2, \quad t \geq 1.
\end{aligned}$$

Here (ii) is due to $a < 1$ and (iii) is because $2R^2 \leq 1 + R^4 \leq d + R^2(\|\mathbf{x}\| + R)^2$. This proves the first inequality of Lemma F.3.

Same as above, When $t < 1$, we know that $f(t) = \Theta(1)$, $f'(t) = \Theta(1)$, $g(t) = \Theta(\sqrt{t})$, $g'(t) = \Theta(\frac{1}{\sqrt{t}})$. Hence by Lemma H.2, we have:

$$\begin{aligned}
\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| &\lesssim \frac{\|\mathbf{x}\| + R}{t^2} + \frac{(\|\mathbf{x}\| + R)^2 R^2}{t^2} + \frac{(\|\mathbf{x}\| + R)^3}{t^3} + \frac{R}{t} \\
&\lesssim \frac{\|\mathbf{x}\| + R}{t^2} + \frac{(\|\mathbf{x}\| + R)^2 R^2}{t^2} + \frac{(\|\mathbf{x}\| + R)^3}{t^3}, \quad t < 1.
\end{aligned}$$

We next consider the case $t \geq 1$. Denote e^{-t} by a , then we know that $f(t) = \Theta(a)$, $f'(t) = \Theta(a)$, $g(t) = \Theta(1)$, $g'(t) = \Theta(a^2)$. By, Lemma H.2, we have:

$$\begin{aligned}
\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| &\lesssim a^2(\|\mathbf{x}\| + aR) + a(\|\mathbf{x}\| + aR)^2 R^2 + a^2(\|\mathbf{x}\| + aR)^3 + aR \\
&\stackrel{(i)}{\lesssim} a(\|\mathbf{x}\| + R) + a(\|\mathbf{x}\| + R)^2 R^2 + a(\|\mathbf{x}\| + R)^3 \\
&\leq (\|\mathbf{x}\| + R) + (\|\mathbf{x}\| + R)^2 R^2 + (\|\mathbf{x}\| + R)^3, \quad t \geq 1.
\end{aligned}$$

Here (i) is due to $a = e^{-t} < 1$ given $t \geq 1$. This completes the proof of Lemma F.3. \square

Proof of Lemma F.4. Because $\mathbf{X}_t = e^{-t} \mathbf{X}_0 + N(0, (1 - e^{-2t}) \mathbf{I}_d)$, we have $f(t) = e^{-t}$ and $g(t) = \sqrt{1 - e^{-2t}}$ in this case. When $t < 1$, we know that

$$f(t) = \Theta(1), f'(t) = \Theta(1), g(t) = \Theta(\sqrt{t}), g'(t) = \Theta(\frac{1}{\sqrt{t}}).$$

Then by Lemma H.1, we can easily show the three inequalities are valid when $t < 1$. Moreover, when $t \geq 1$, we have

$$f(t) = \Theta(e^{-t}), f'(t) = \Theta(e^{-t}), g(t) = \Theta(1), g'(t) = \Theta(e^{-2t}).$$

Then by Lemma H.1, we can easily show the three inequalities are valid when $t \geq 1$. This completes the proof of Lemma F.4. \square

H.3 PROOF OF LEMMAS G.3 AND G.4

Proof of Lemma G.3. Since $\mathbf{X}_t = \mathbf{X}_0 + N(0, \sqrt{t} \mathbf{I}_d)$, we have $f(t) = 1$ and $g(t) = \sqrt{t}$ in this case. Hence $f'(t) = 0$ and $g'(t) = \Theta(\frac{1}{\sqrt{t}})$. By Lemma H.2, we have:

$$\left\| \frac{\partial}{\partial t} (\nabla \log q(t, \mathbf{x})) \right\| \lesssim \frac{\|\mathbf{x}\| + R}{t^2} + \frac{(\|\mathbf{x}\| + R)^3}{t^3}.$$

$$\begin{aligned}
\left| \frac{\partial}{\partial t} (\text{tr}(\nabla^2 \log q(t, \mathbf{x}))) \right| &\lesssim \frac{d}{t^2} + \frac{R^2}{t^3} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4} \\
&\stackrel{(i)}{\lesssim} \frac{d}{t^2} + \frac{R^2(\|\mathbf{x}\| + R)^2}{t^4},
\end{aligned}$$

here (i) is due to $2\frac{R^2}{t^3} \leq \frac{1}{t^2} + \frac{R^4}{t^4}$. This completes our proof. \square

Proof of Lemma G.4. Since $\mathbf{X}_t = \mathbf{X}_0 + N(0, \sqrt{t} \mathbf{I}_d)$, we have $f(t) = 1$ and $g(t) = \sqrt{t}$ in this case. Hence $f'(t) = 0$ and $g'(t) = \Theta(\frac{1}{\sqrt{t}})$. Substituting into Lemma H.1, we completes our proof. \square

2484 **I PROOF OF REMAINING LEMMAS IN SECTIONS F AND G**
 2485

2486 **I.1 PROOF OF LEMMAS F.1 AND F.2**

2487 We first present the following two technical lemmas.

2488 **Lemma I.1.** Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (F.2) and (F.3). Then for any $0 \leq k \leq N - 1$,
 2489 $t \in [t_k, t_{k+1}]$ and $\mathbf{x} \in \mathbb{R}^d$, we have:

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq \frac{e^{|\nabla F_t(F_t^{-1}(\mathbf{x}))|}}{a} \cdot \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2} \cdot e^{\frac{(1+d/2)\|\mathbf{g}(\mathbf{x})\|_2^2}{2(a^2-1)}},$$

2494 where $a = e^{t-t_k}$ and $\mathbf{g}(\mathbf{x}) = aF_t^{-1}(\mathbf{x}) - \mathbf{x}$.

2495 **Lemma I.2.** Suppose $\mathbf{Y}_t = e^{-(T-t)}\mathbf{X}_{\text{data}} + N(0, (1 - e^{-2(T-t)})I_d)$ and $p(\|\mathbf{X}_{\text{data}}\|_2 < R) = 1$.
 2496 Then for $\lambda < \min\{\frac{1}{d}, \frac{1}{2}\}$, we have:

$$\mathbb{E}e^{\lambda\|\mathbf{Y}_t\|_2^2} \leq e^{\lambda R^2 - 1}.$$

2500 Using these lemmas, we can start the proof of Lemma F.1:

2501 *Proof of Lemma F.1.* To start with, we have the following lemma about the ratio: Firstly, using the
 2502 expression of the interpolation operator (F.3), we have

$$\nabla F_t(\mathbf{x}) = e^{t-t_k} \mathbf{I}_d + (e^{t-t_k} - 1) \nabla s_\theta(T - t_k, \mathbf{x}).$$

2505 The operator norm of $\nabla F_t(\mathbf{x})$ can be bounded by

$$\begin{aligned} \|\nabla F_t(\mathbf{x})\|_2 &\leq e^{t-t_k} + (e^{t-t_k} - 1)L \\ &= a + (a-1)L \end{aligned}$$

2510 Using the fact that $|\mathbf{A}| \leq \|\mathbf{A}\|^d$, we know that:

$$\begin{aligned} |\nabla F_t(F_t^{-1}(\mathbf{x}))| &\leq (a + (a-1)L)^d \\ &\leq (1 + 2\eta + 2\eta L)^d \\ &\leq e \end{aligned}$$

2516 Here the last inequality is because $\eta \leq \frac{1}{2(L+1)d}$. With our time schedule, by (2.6), we have

$$t - t_k \leq t_{k+1} - t_k \leq \eta \min\{1, T - t\}.$$

2519 We consider two cases when $T - t \geq 1$ and $T - t < 1$.

2520 First, when $T - t \geq 1$, we have $t - t_k \leq \eta$. Therefore, we have

$$\begin{aligned} \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2} &\leq \left(1 + \frac{e^{2(t-t_k)} - 1}{1 - e^{-2(T-t)}} \right)^{d/2} \\ &\leq \left(1 + \frac{e^{2\eta} - 1}{1 - e^{-2}} \right)^{d/2}. \end{aligned}$$

2527 Since $e^x - 1 \leq 2x$ when $x \leq 1$ and $1 - e^{-2} > \frac{1}{2}$, when $\eta \leq 1/(8d)$, we have:

$$\left(1 + \frac{e^{2\eta} - 1}{1 - e^{-2}} \right)^{d/2} \leq \left(1 + \frac{1}{d} \right)^{d/2} \leq \sqrt{e}.$$

2531 Secondly, when $T - t < 1$, since $1 - e^{-2x} > x/4$ and $e^x - 1 \leq 2x$ when $0 < x < 1$, when
 2532 $\eta \leq 1/16d$, we know that:

$$\begin{aligned} \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2} &\leq \left(1 + 4 \frac{e^{2(t-t_k)} - 1}{T - t} \right)^{d/2} \\ &\leq \left(1 + 4 \frac{4(t - t_k)}{T - t} \right)^{d/2} \end{aligned}$$

$$2538 \quad \leq (1 + 16\eta)^{d/2} \leq \sqrt{e}. \\ 2539$$

2540 Combining the two cases, we can summarize that when $\eta \leq 1/16d$, we have

$$2541 \quad \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2} \leq \sqrt{e}. \\ 2542 \\ 2543$$

2544 Since $g(\mathbf{x}) = aF_t^{-1}(\mathbf{x}) - \mathbf{x}$, we have:

$$2545 \quad g(\mathbf{x}) = aF_t^{-1}(\mathbf{x}) - \mathbf{x} \\ 2546 \quad = aF_t^{-1}(\mathbf{x}) - aF_t^{-1}(\mathbf{x}) - (a-1)\mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x})) \\ 2547 \quad = -(a-1)\mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x})). \\ 2548 \\ 2549$$

2550 Hence,

$$2551 \quad \frac{(1+d/2)\|\mathbf{g}(\mathbf{x})\|_2^2}{2(a^2-1)} = \frac{(1+d/2)(a-1)\|\mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\|^2}{2(a+1)}. \\ 2552 \\ 2553$$

2554 Using the assumption on \mathbf{s}_θ and inequality (F.4), we have that:

$$2555 \quad \frac{(1+d/2)(a-1)\|\mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\|^2}{2(a+1)} \leq \frac{(1+d/2)(a-1)}{2(a+1)}(L\|F_t^{-1}(\mathbf{x})\| + c)^2 \\ 2556 \\ 2557 \quad \leq \frac{(1+d/2)(a-1)}{2(a+1)}(2L\|\mathbf{x}\| + L + c)^2 \\ 2558 \\ 2559 \quad \leq \frac{(1+d/2)(a-1)}{2(a+1)} \\ 2560 \\ 2561 \quad \leq \frac{(1+d/2)(a-1)}{2(a+1)} \\ 2562 \\ 2563 \quad \leq \frac{(1+d/2)(a-1)}{a+1}4L^2\|\mathbf{x}\|^2 + \frac{(1+d/2)(a-1)}{a+1}(L+c)^2 \\ 2564 \\ 2565 \quad \leq (t-t_k) \underbrace{(1+d/2)4L^2\|\mathbf{x}\|^2}_{c_1} + (t-t_k) \underbrace{(1+d/2)(L+c)^2}_{c_2}, \\ 2566$$

2567 where the 1st inequality holds due to $a \geq 1$ and $a-1 \leq 2(t-t_k)$. Thus, we know that:

$$2568 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2}. \\ 2569 \\ 2570$$

2571 This proves the first part of Lemma F.1.

2572 Next we prove the second part, we know that:

$$2573 \quad \int_{\Omega_t} \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim \int_{\Omega_t} e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ 2574 \\ 2575 \quad \leq e^{(t-t_k)c_2} \mathbb{E} e^{(t-t_k)c_1\|\mathbf{Y}_t\|^2} \\ 2576 \\ 2577 \quad \stackrel{(i)}{\leq} e^{(t-t_k)c_2} e^{(t-t_k)c_1R^2-1} \\ 2578 \\ 2579 \quad \stackrel{(ii)}{\lesssim} 1, \\ 2580 \\ 2581$$

2582 where the (i) holds due to Lemma I.2 since our $t-t_k \leq \eta \leq \min\{\frac{1}{c_1d}, \frac{1}{2c_1}\}$. And (ii) holds due to
2583 $t-t_k \leq \eta \leq \min\{\frac{1}{c_2}, \frac{1}{R^2c_1}\}$

2584 By Lemma I.1, we know that:

$$2585 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq \frac{e^{|\nabla F_t(F_t^{-1}(\mathbf{x}))|}}{a} \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2} \cdot e^{\frac{(1+d/2)\|\mathbf{g}(\mathbf{x})\|_2^2}{2(a^2-1)}} \\ 2586 \\ 2587 \\ 2588 \\ 2589 \lesssim 1.$$

2590 here $a = e^{t-t_k}$ and $\mathbf{g}(\mathbf{x}) = aF_t^{-1}(\mathbf{x}) - \mathbf{x}$. This completes the proof of Lemma F.1. \square

2591 Next, we begin our proof of Lemma F.2.

Proof of Lemma F.2. To start with, using Lemma F.1, we know that:

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2},$$

where $c_1 = (1+d/2)4L^2$ and $c_2 = (1+d/2)(L+c)^2$.

Thus, we have:

$$\begin{aligned} \int_{\Omega_t} \left\| \frac{\nabla q(T-t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} &= \int_{\Omega_t} \left\| \nabla \log q(T-t, \mathbf{x}) \right\|^2 \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ &\lesssim \int_{\Omega_t} \left\| \nabla \log q(T-t, \mathbf{x}) \right\|^2 e^{\eta(c_1\|\mathbf{x}\|^2 + c_2)} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ &\lesssim \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} \right], \end{aligned}$$

where the last inequality holds due to $\eta \leq \frac{1}{c_2}$. Next, we select a constant $M = 2R$. We have the following inequality:

$$\begin{aligned} &\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} \right] \\ &= \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| < M} \right] + \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right] \\ &\lesssim \underbrace{\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 \right]}_{I_1} + \underbrace{\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right]}_{I_2}, \end{aligned}$$

where the last inequality holds due to $\eta \leq \frac{1}{4R^2 c_1} = \frac{1}{c_1 M^2}$. For I_1 , using Lemma K.2, we have:

$$\begin{aligned} I_1 &= \mathbb{E}_Q \left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 \\ &\leq d \frac{1}{1 - e^{-2(T-t)}} \\ &\stackrel{(i)}{\lesssim} \frac{d}{\min\{T-t, 1\}}, \end{aligned} \tag{I.1}$$

where (i) holds because $1 - e^{-2x} > x/2$ when $x < 1$ and $1 - e^{-2x} > 1/2$ when $x \geq 1$.

For I_2 , using Lemma H.1 with $f(t) = e^{-(T-t)}$, $g(t) = \sqrt{1 - e^{-2(T-t)}}$, we have:

$$\begin{aligned} \left\| \nabla \log q(T-t, \mathbf{x}) \right\| &\leq \frac{\|\mathbf{x}\| + R}{1 - e^{-2(T-t)}} \\ &= \frac{\|\mathbf{x}\| + R}{\sigma_{T-t}}, \end{aligned}$$

here we denote $1 - e^{-2(T-t)}$ by σ_{T-t} . Therefore, we have:

$$I_2 \lesssim \mathbb{E}_Q \left[\left(\frac{\|\mathbf{Y}_t\| + R}{\sigma_{T-t}} \right)^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right]. \tag{I.2}$$

Let $\alpha = e^{-(T-t)}$, we have $\mathbf{Y}_t = \mathbf{X}_{T-t} = \alpha \mathbf{X}_0 + \sqrt{1-\alpha^2} \mathbf{Z}$. Thus, $\|\mathbf{Y}_t\| \leq R + \sqrt{1-\alpha^2} \|\mathbf{Z}\|$. Hence, (I.2) becomes

$$\begin{aligned} I_2 &\leq \frac{1}{(\sigma_{T-t})^2} \mathbb{E}_Q \left[(2R + \sqrt{1-\alpha^2} \|\mathbf{Z}\|)^2 e^{\eta c_1 (R + \sqrt{1-\alpha^2} \|\mathbf{Z}\|)^2} 1_{\|\mathbf{Z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} \right] \\ &\lesssim \frac{1}{(\sigma_{T-t})^2} \mathbb{E}_Q \left[(R^2 + (1-\alpha^2) \|\mathbf{Z}\|^2) e^{2\eta c_1 (1-\alpha^2) \|\mathbf{Z}\|^2} 1_{\|\mathbf{Z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} \right], \end{aligned} \tag{I.3}$$

where the last inequality holds due to $(a+b)^2 \leq 2a^2 + 2b^2$, and $\eta \leq 1/c_1 R^2$. Let $\lambda = 2\eta c_1 (1-\alpha^2)$. Then we have

$$I_2 \lesssim \frac{R^2}{(\sigma_{T-t})^2} \underbrace{\int_{\mathbb{R}^d} \frac{1}{(\sqrt{2\pi})^d} e^{-\frac{\|\mathbf{z}\|^2}{2}} e^{\lambda \|\mathbf{z}\|^2} 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} d\mathbf{z}}_{K_1}$$

$$+ \frac{2}{\sigma_{T-t}} \underbrace{\int_{\mathbb{R}^d} \frac{1}{\sqrt{2\pi}^d} e^{-\frac{\|\mathbf{z}\|^2}{2}} \|\mathbf{z}\|^2 e^{\lambda \|\mathbf{z}\|^2} 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} d\mathbf{z},}_{K_2} \quad (I.4)$$

here we use that $1 - \alpha^2 = 1 - e^{-2(T-t)} = \sigma_{T-t}$. Let $\phi_{\sigma^2}(\mathbf{x}) = \exp(-\|\mathbf{x}\|^2/2\sigma^2)/(2\pi\sigma^2)^{d/2}$. When λ smaller than $\frac{1}{2}$, we have $e^{\lambda \|\mathbf{z}\|^2} \phi_1(\mathbf{z}) = \phi_{\frac{1}{1-2\lambda}}(\mathbf{z})(\frac{1}{1-2\lambda})^{\frac{d}{2}}$. Therefore, we know that:

$$\begin{aligned} K_1 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \int_{\mathbb{R}^d} \phi_{\frac{1}{1-2\lambda}}(\mathbf{z}) 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} d\mathbf{z} \\ &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{P}\left[\|\mathbf{Z}'\| \geq \frac{M-R}{\sqrt{1-\alpha^2}} \mid \mathbf{Z}' \sim N\left(0, \frac{1}{1-2\lambda} \mathbf{I}_d\right)\right] \\ &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{P}\left[\|\mathbf{Z}\| \geq \frac{\sqrt{1-2\lambda}(M-R)}{\sqrt{1-\alpha^2}} \mid \mathbf{Z} \sim N(0, \mathbf{I}_d)\right] \\ &\stackrel{(i)}{\leq} \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \frac{1-\alpha^2}{(1-2\lambda)(M-R)^2} \mathbb{E}\|\mathbf{Z}\|^2 \\ &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}+1} \frac{\sigma_{T-t} d}{(M-R)^2} \\ &\stackrel{(i)}{\lesssim} \frac{\sigma_{T-t} d}{R^2}. \end{aligned} \quad (I.5)$$

where (i) holds due to the Markov's inequality. Since $\eta \leq \frac{1}{2c_1 d} \leq \frac{1}{2\sigma_{T-t} c_1 d}$, we have $\lambda = 2\eta c_1 (1 - \alpha^2) \leq 1/(d+2)$. This implies $(1/(1-2\lambda))^{d/2} \leq e$, thus (ii) holds. Moreover, for K_2 we have:

$$\begin{aligned} K_2 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \int_{\mathbb{R}^d} \phi_{\frac{1}{1-2\lambda}}(\mathbf{z}) \|\mathbf{z}\|^2 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{1-\alpha^2}}} d\mathbf{z} \\ &\leq \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{E}_{\mathbf{Z}' \sim N(0, \frac{1}{1-2\lambda} \mathbf{I}_d)} \|\mathbf{Z}'\|^2 \\ &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \frac{d}{1-2\lambda} \\ &\lesssim d. \end{aligned} \quad (I.6)$$

The last inequality holds due to $(1/(1-2\lambda))^{d/2} \leq e$ when $\lambda \leq 1/(d+2)$. Substituting (I.5) and (I.6) into (I.2), we have

$$\begin{aligned} I_2 &\lesssim \frac{d}{\sigma_{T-t}} \\ &\lesssim \frac{d}{\min\{T-t, 1\}}. \end{aligned} \quad (I.7)$$

Combining I.1 and I.7, we have:

$$\int_{\Omega_t} \left\| \frac{\nabla q(T-t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim \frac{d}{\min\{T-t, 1\}},$$

which completes the proof of Lemma F.2. \square

I.2 PROOF OF LEMMAS G.1 AND G.2

We first present three technical lemmas.

Lemma I.3. Let \mathbf{Y}_t and $F_t(\mathbf{z})$ be defined in (G.2) and (G.5). For any $k, t \in [t_k, t_{k+1}]$ and $\mathbf{x} \in \mathbb{R}^d$, we have:

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq e \left| \nabla F_t(F_t^{-1}(\mathbf{x})) \right| \cdot \left(\frac{T-t + \frac{t-t_k}{1+2/d}}{T-t} \right)^{d/2} \cdot e^{\frac{(1+d/2)\|F_t^{-1}(\mathbf{x}) - \mathbf{x}\|_2^2}{2(t-t_k)}},$$

Lemma I.4. Suppose $\mathbf{Y}_t = \mathbf{X}_0 + N(0, (T-t)\mathbf{I}_d)$ and $p(\|\mathbf{X}_0\|_2 < R) = 1$. Then for $\lambda < \min\{\frac{1}{4d(T-t)}, \frac{1}{R^2}\}$, we have:

$$\mathbb{E} e^{\lambda \|\mathbf{Y}_t\|_2^2} \lesssim 1$$

2700 **Lemma I.5.** Recall that \mathbf{X}_t is our forward process defined in G.1 and we use $q(t, \mathbf{x})$ to denote its
 2701 law, under Assumption 5.1, we have:

$$2702 \quad \mathbb{E} \|\nabla \log q(t, \mathbf{X}_t)\|^2 \leq \frac{d}{t}.$$

2705 *Proof of lemma G.1.* By Lemma I.3, we know that:

$$2706 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq e^{\left| \nabla F_t(F_t^{-1}(\mathbf{x})) \right| \cdot e^{\frac{(1+d/2)\|F_t^{-1}(\mathbf{x}) - \mathbf{x}\|_2^2}{2(t-t_k)}} \cdot \left(\frac{T-t+\frac{t-t_k}{1+2/d}}{T-t} \right)^{d/2}},$$

2709 Firstly, using the expression of the interpolation operator (G.5), we have

$$2710 \quad \nabla F_t(\mathbf{x}) = \mathbf{I}_d + (t-t_k)c_l \nabla \mathbf{s}_\theta(T-t_k, \mathbf{x}).$$

2711 The operator norm of $\nabla F_t(\mathbf{x})$ can be bounded by

$$2713 \quad \begin{aligned} \|\nabla F_t(\mathbf{x})\|_2 &\leq 1 + (t-t_k)c_l L \\ 2714 &\leq 1 + (t-t_k)L, \end{aligned}$$

2715 Using the fact that $|\mathbf{A}| \leq \|\mathbf{A}\|^d$, we know that:

$$2717 \quad |\nabla F_t(F_t^{-1}(\mathbf{x}))| \leq (1 + (t-t_k)L)^d.$$

2718 Since $t-t_k \leq \eta < 1$, when $\eta \leq 1/(Ld)$, we have

$$2720 \quad |\nabla F_t(F_t^{-1}(\mathbf{x}))| \leq (1 + \eta L)^d \leq e.$$

2721 With our time schedule, by (2.6), we have

$$2722 \quad t-t_k \leq t_{k+1}-t_k \leq \eta \min\{1, T-t\}.$$

2723 We have $t-t_k \leq \eta$. Therefore, we have

$$2725 \quad \begin{aligned} \left(\frac{T-t+\frac{t-t_k}{1+2/d}}{T-t} \right)^{d/2} &\leq \left(1 + \frac{t-t_k}{T-t} \right)^{d/2} \\ 2727 &\leq (1+\eta)^{d/2}. \end{aligned}$$

2729 When $\eta \leq \frac{1}{d}$, we have:

$$2731 \quad (1+\eta)^{d/2} \leq (1+\frac{1}{d})^{d/2} \leq \sqrt{e}.$$

2732 Then we have:

$$2733 \quad \begin{aligned} \frac{(1+d/2)\|F_t^{-1}(\mathbf{x}) - \mathbf{x}\|_2^2}{2(t-t_k)} &= \frac{(1+d/2)(t-t_k)}{2} c_l^2 \|\mathbf{s}_\theta(T-t_k, F_t^{-1}(\mathbf{x}))\|^2 \\ 2734 &\leq \frac{(1+d/2)(t-t_k)}{2} (L\|F_t^{-1}(\mathbf{x})\| + c)^2 \\ 2736 &\leq \frac{(1+d/2)(t-t_k)}{2} (2L\|\mathbf{x}\| + L + c)^2 \\ 2738 &\leq (t-t_k) \left(\underbrace{(1+d/2)4L^2\|\mathbf{x}\|^2}_{c_1} + \underbrace{(1+d/2)(L+c)^2}_{c_2} \right). \end{aligned}$$

2742 Putting together, we know that:

$$2744 \quad \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2 + (t-t_k)c_2}.$$

2746 Moreover, since $t-t_k \leq \eta \leq 1/c_2$, we have:

$$2748 \quad \begin{aligned} \int_{\Omega_t} \left(\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right)^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} &= \int_{\Omega_t} \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ 2750 &\lesssim \mathbb{E} e^{(t-t_k)c_1\|\mathbf{Y}_t\|^2} \\ 2751 &\stackrel{(i)}{\lesssim} 1, \end{aligned}$$

2753 where (i) holds due to Lemma I.4 (we have $(t-t_k)c_1 \leq \min\{\frac{1}{4d(T-t)}, \frac{1}{R^2}\}$). \square

Proof of Lemma G.2. To start with, recall that we assume that $L \geq 1$, we can easily verify that η satisfies Lemma G.1's condition. Using Lemma G.1, we know that:

$$\begin{aligned} \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} &\lesssim e^{(t-t_k)c_1\|\mathbf{x}\|^2+(t-t_k)c_2} \\ &\leq e^{\eta(c_1\|\mathbf{x}\|^2+c_2)}, \end{aligned}$$

where $c_1 = (1+d/2)4L^2$ and $c_2 = (1+d/2)(L+c)^2$. Thus we have:

$$\begin{aligned} \int_{\Omega_t} \left\| \frac{\nabla q(T-t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} &\leq \int_{\Omega_t} \left\| \nabla \log q(T-t, \mathbf{x}) \right\|^2 \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ &\lesssim \int_{\Omega_t} \left\| \nabla \log q(T-t, \mathbf{x}) \right\|^2 e^{\eta(c_1\|\mathbf{x}\|^2+c_2)} p_{\mathbf{Y}_t}(\mathbf{x}) d\mathbf{x} \\ &\lesssim \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} \right], \end{aligned}$$

where the last inequality holds due to $\eta \leq \frac{1}{c_2}$. Next, we select a constant $M = 2R$. We have the following inequality:

$$\begin{aligned} &\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} \right] \\ &= \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| < M} \right] + \mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right] \\ &\lesssim \underbrace{\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 \right]}_{I_1} + \underbrace{\mathbb{E}_Q \left[\left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right]}_{I_2}, \end{aligned}$$

where the last inequality holds due to $\eta \leq \frac{1}{c_1 R^2}$. For I_1 , using Lemma I.5, we have:

$$\begin{aligned} I_1 &= \mathbb{E}_Q \left\| \nabla \log q(T-t, \mathbf{Y}_t) \right\|^2 \\ &\leq \frac{d}{T-t}, \end{aligned} \tag{I.8}$$

For I_2 , using Lemma H.1 with $f(t) = 1$, $g(t) = \sqrt{T-t}$, we have:

$$\left\| \nabla \log q(T-t, \mathbf{x}) \right\| \leq \frac{\|\mathbf{x}\| + R}{T-t},$$

Therefore, we have:

$$I_2 \leq \mathbb{E}_Q \left[\left(\frac{\|\mathbf{Y}_t\| + R}{T-t} \right)^2 e^{\eta c_1 \|\mathbf{Y}_t\|^2} 1_{\|\mathbf{Y}_t\| \geq M} \right]. \tag{I.9}$$

We have $\mathbf{Y}_t = \mathbf{X}_{T-t} = \mathbf{X}_0 + \sqrt{T-t}\mathbf{Z}$. Thus, $\|\mathbf{Y}_t\| \leq R + \sqrt{T-t}\|\mathbf{Z}\|$. Hence, (I.9) becomes

$$\begin{aligned} I_2 &\leq \frac{1}{(T-t)^2} \mathbb{E}_Q \left[(2R + \sqrt{T-t}\|\mathbf{Z}\|)^2 e^{\eta c_1 (R + \sqrt{T-t}\|\mathbf{Z}\|)^2} 1_{\|\mathbf{Z}\| \geq \frac{M-R}{\sqrt{T-t}}} \right] \\ &\lesssim \frac{1}{(T-t)^2} \mathbb{E}_Q \left[(R^2 + (T-t)\|\mathbf{Z}\|^2) e^{2\eta c_1 (T-t)\|\mathbf{Z}\|^2} 1_{\|\mathbf{Z}\| \geq \frac{M-R}{\sqrt{T-t}}} \right], \end{aligned} \tag{I.10}$$

where the last inequality holds due to $(a+b)^2 \leq 2a^2 + 2b^2$, and $\eta \leq \frac{1}{c_1 R^2}$. Let $\lambda = 2\eta c_1(T-t)$. Then we have

$$\begin{aligned} I_2 &\lesssim \frac{R^2}{(T-t)^2} \underbrace{\int_{\mathbb{R}^d} \frac{1}{(\sqrt{2\pi})^d} e^{-\frac{\|\mathbf{z}\|^2}{2}} e^{\lambda \|\mathbf{z}\|^2} 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{T-t}}} d\mathbf{z}}_{K_1} \\ &\quad + \frac{2}{T-t} \underbrace{\int_{\mathbb{R}^d} \frac{1}{\sqrt{2\pi}^d} e^{-\frac{\|\mathbf{z}\|^2}{2}} \|\mathbf{z}\|^2 e^{\lambda \|\mathbf{z}\|^2} 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{T-t}}} d\mathbf{z}}_{K_2}, \end{aligned} \tag{I.11}$$

Let $\phi_{\sigma^2}(\mathbf{x}) = \exp(-\|\mathbf{x}\|^2/2\sigma^2)/(2\pi\sigma^2)^{d/2}$. When λ smaller than $\frac{1}{2}$, we have $e^{\lambda\|\mathbf{z}\|^2}\phi_1(\mathbf{z}) = \phi_{\frac{1}{1-2\lambda}}(\mathbf{z})(\frac{1}{1-2\lambda})^{\frac{d}{2}}$. Therefore, we know that:

$$\begin{aligned}
 K_1 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \int_{\mathbb{R}^d} \phi_{\frac{1}{1-2\lambda}}(\mathbf{z}) 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{T-t}}} d\mathbf{z} \\
 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{P}\left[\|\mathbf{Z}'\| \geq \frac{M-R}{\sqrt{T-t}} \mid \mathbf{Z}' \sim N\left(0, \frac{1}{1-2\lambda} \mathbf{I}_d\right)\right] \\
 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{P}\left[\|\mathbf{Z}\| \geq \frac{\sqrt{1-2\lambda}(M-R)}{\sqrt{T-t}} \mid \mathbf{Z} \sim N(0, \mathbf{I}_d)\right] \\
 &\stackrel{(i)}{\leq} \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \frac{T-t}{(1-2\lambda)(M-R)^2} \mathbb{E}\|\mathbf{Z}\|^2 \\
 &\leq \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}+1} \frac{(T-t)d}{(M-R)^2} \\
 &\stackrel{(i)}{\lesssim} \frac{(T-t)d}{R^2}. \tag{I.12}
 \end{aligned}$$

where (i) holds due to the Markov's inequality. (ii) holds because $\eta \leq \frac{1}{2(d+2)c_1(T-t)}$, thus $\lambda \leq 1/(d+2)$, then we know that $(1/(1-2\lambda))^{d/2} \leq e$. Moreover, for K_2 we have:

$$\begin{aligned}
 K_2 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \int_{\mathbb{R}^d} \phi_{\frac{1}{1-2\lambda}}(\mathbf{z}) \|\mathbf{z}\|^2 1_{\|\mathbf{z}\| \geq \frac{M-R}{\sqrt{T-t}}} d\mathbf{z} \\
 &\leq \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \mathbb{E}_{\mathbf{Z}' \sim N(0, \frac{1}{1-2\lambda} \mathbf{I}_d)} \|\mathbf{Z}'\|^2 \\
 &= \left(\frac{1}{1-2\lambda}\right)^{\frac{d}{2}} \frac{d}{1-2\lambda} \\
 &\lesssim d. \tag{I.13}
 \end{aligned}$$

The last inequality holds due to $(1/(1-2\lambda))^{d/2} \leq e$. Substituting (I.12) and (I.13) into (I.11), we have

$$I_2 \lesssim \frac{d}{T-t}. \tag{I.14}$$

Combining I.8 and I.14, we have:

$$\int_{\Omega_t} \left\| \frac{\nabla q(T-t, \mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \right\|^2 p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) d\mathbf{x} \lesssim \frac{d}{T-t},$$

which completes the proof of Lemma G.2. \square

J PROOF OF LEMMAS IN SECTION I

Proof of Lemma I.1. Using the Jacobian transformation of probability densities, we have

$$p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) = p_{\mathbf{Y}_{t_k}}(F_t^{-1}(\mathbf{x})) |\nabla(F_t^{-1}(\mathbf{x}))|. \tag{J.1}$$

Moreover, the forward process indicates $\mathbf{Y}_{t_k} = e^{-(t-t_k)} \mathbf{Y}_t + \mathbf{Z}_{t,t_k}$, where $\mathbf{Z}_{t,t_k} \sim N(0, 1 - e^{-2(t-t_k)})$ is independent of \mathbf{Y}_t . Using the Jacobian transformation of probability densities, we have

$$\begin{aligned}
 p_{\mathbf{Y}_{t_k}}(F_t^{-1}(\mathbf{x})) &= p_{\mathbf{Y}_t + e^{t-t_k} \mathbf{Z}_{t,t_k}}(e^{t-t_k} F_t^{-1}(\mathbf{x})) \cdot e^{t-t_k} \\
 &= e^{t-t_k} \int_{\mathbb{R}^d} q(t, \mathbf{y}) \phi_{e^{2(t-t_k)}-1}(e^{t-t_k} F_t^{-1}(\mathbf{x}) - \mathbf{y}) d\mathbf{y},
 \end{aligned}$$

where $\phi_{\sigma^2}(\mathbf{x}) = \exp(-\|\mathbf{x}\|^2/2\sigma^2)/(2\pi\sigma^2)^{d/2}$ is the probability density function of Gaussian distribution with variance σ^2 . The last equality holds due to the formula for the sum of independent

variables. For simplicity, denote $a = e^{t-t_k}$ and we ignore the dependency of t and t_k when it will not cause any confusion. Then we have

$$p_{\mathbf{Y}_{t_k}}(F_t^{-1}(\mathbf{x})) = a \int_{\mathbb{R}^d} q(T-t, \mathbf{y}) \phi_{a^2-1}(aF_t^{-1}(\mathbf{x}) - \mathbf{y}) d\mathbf{y}. \quad (\text{J.2})$$

Since $aF_t^{-1}(\mathbf{x}) - \mathbf{y} = \mathbf{x} - \mathbf{y} + (aF_t^{-1}(\mathbf{x}) - \mathbf{x})$, with the shorthand notation $g(\mathbf{x}) = aF_t^{-1}(\mathbf{x}) - \mathbf{x}$, we have

$$\begin{aligned} \phi_{a^2-1}(aF_t^{-1}(\mathbf{x}) - \mathbf{y}) &= \phi_{a^2-1}(\mathbf{x} - \mathbf{y} + g(\mathbf{x})) \\ &= \frac{1}{(2\pi(a^2-1))^{d/2}} e^{-\frac{\|\mathbf{x}-\mathbf{y}+g(\mathbf{x})\|_2^2}{2(a^2-1)}} \\ &= \frac{1}{(2\pi(a^2-1))^{d/2}} e^{-\frac{\|\mathbf{x}-\mathbf{y}\|_2^2 + \|g(\mathbf{x})\|_2^2 + 2\langle \mathbf{x}-\mathbf{y}, g(\mathbf{x}) \rangle}{2(a^2-1)}} \\ &\geq \frac{1}{(2\pi(a^2-1))^{d/2}} e^{-\frac{(1+2/d)\|\mathbf{x}-\mathbf{y}\|_2^2 + (1+d/2)\|g(\mathbf{x})\|_2^2}{2(a^2-1)}} \\ &= \frac{1}{(1+2/d)^{d/2}} \cdot \phi_{\frac{a^2-1}{1+2/d}}(\mathbf{x} - \mathbf{y}) \cdot e^{-\frac{(1+d/2)\|g(\mathbf{x})\|_2^2}{2(a^2-1)}}, \end{aligned} \quad (\text{J.3})$$

where the first inequality holds due to the Young's inequality. In the last equality, we use the definition of $\phi_{\sigma^2}(\cdot)$ with $\sigma^2 = (a^2-1)/(1+2/d)$. Combining (J.1), (J.2) and (J.3), we have

$$\begin{aligned} p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) &\geq |\nabla(F_t^{-1}(\mathbf{x}))| \cdot a \cdot \int_{\mathbb{R}^d} q(t, \mathbf{y}) \frac{1}{(1+2/d)^{d/2}} \phi_{\frac{a^2-1}{1+2/d}}(\mathbf{x} - \mathbf{y}) e^{-\frac{(1+d/2)\|g(\mathbf{x})\|_2^2}{2(a^2-1)}} d\mathbf{y} \\ &\geq \frac{a}{e |\nabla F_t(F_t^{-1}(\mathbf{x}))|} p_{\mathbf{Y}_t + N(0, \frac{a^2-1}{1+2/d})}(\mathbf{x}) \cdot e^{-\frac{(1+d/2)\|g(\mathbf{x})\|_2^2}{2(a^2-1)}}, \end{aligned}$$

where we use $(1+2/d)^{d/2} \leq e$, the fact that $\nabla(F_t^{-1}(\mathbf{x})) = [\nabla F_t(F_t^{-1}(\mathbf{x}))]^{-1}$ and the formula for the distribution of the sum of independent random variables. Then we have

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq \frac{e |\nabla F_t(F_t^{-1}(\mathbf{x}))|}{a} \cdot e^{\frac{(1+d/2)\|g(\mathbf{x})\|_2^2}{2(a^2-1)}} \cdot \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{\mathbf{Y}_t + N(0, \frac{a^2-1}{1+2/d})}(\mathbf{x})}$$

Since $\mathbf{Y}_t = e^{-(T-t)} X_{\text{data}} + N(0, 1 - e^{-2(T-t)})$, by Lemma K.3, we know that

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{\mathbf{Y}_t + N(0, \frac{t-t_k}{1+2/d})}(\mathbf{x})} \leq \left(\frac{1 - e^{-2(T-t)} + \frac{a^2-1}{1+2/d}}{1 - e^{-2(T-t)}} \right)^{d/2}$$

Combining the above two inequalities and we can complete the proof of Lemma I.1 \square

Proof of lemma I.2. Since $\mathbf{Y}_t = e^{-(T-t)} X_{\text{data}} + N(0, (1 - e^{-2(T-t)}) I_d)$, denote $e^{-(T-t)}$ by c and use \mathbf{Z} to represent standard normal distribution, we know that:

$$\begin{aligned} \mathbb{E} e^{\lambda \|\mathbf{Y}_t\|_2^2} &= \mathbb{E} e^{\lambda \|cX_{\text{data}} + \sqrt{1-c^2}\mathbf{Z}\|_2^2} \\ &\stackrel{(i)}{\leq} \mathbb{E} e^{\lambda \|X_{\text{data}}\|_2^2 + \lambda \|\mathbf{Z}\|_2^2} \\ &\stackrel{(ii)}{=} \mathbb{E} e^{\lambda \|X_{\text{data}}\|_2^2} \cdot \mathbb{E} e^{\lambda \|\mathbf{Z}\|_2^2}. \end{aligned}$$

where (i) holds due to the Cauchy-Schwartz inequality. (ii) holds due to the independency of X_{data} and \mathbf{Z} .

Since $p(\|X_{\text{data}}\|_2 < R) = 1$, then $\mathbb{E} e^{\lambda \|X_{\text{data}}\|_2^2} \leq e^{\lambda R^2}$. For $\lambda < \frac{1}{2}$, we know that:

$$\begin{aligned} \mathbb{E} e^{\lambda \|\mathbf{Z}\|_2^2} &= \int_{\mathbb{R}^d} e^{\lambda \|\mathbf{x}\|_2^2} (2\pi)^{-\frac{d}{2}} e^{-\frac{\|\mathbf{x}\|_2^2}{2}} d\mathbf{x} \\ &= \int_{\mathbb{R}^d} (2\pi \frac{1}{1-2\lambda})^{-\frac{d}{2}} (1-2\lambda)^{\frac{d}{2}} e^{-\frac{\|\mathbf{x}\|_2^2}{2(1-2\lambda)}} d\mathbf{x} \end{aligned}$$

$$= (1 - 2\lambda)^{\frac{d}{2}}.$$

When $\lambda < \frac{1}{d}$, $(1 - 2\lambda)^{-\frac{d}{2}} < \frac{1}{e}$, thus we have:

$$\mathbb{E} e^{\lambda \|\mathbf{Y}_t\|_2^2} \leq e^{\lambda R^2 - 1}.$$

□

Proof of Lemma I.3. Using the Jacobian transformation of probability densities, we have

$$p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) = p_{\mathbf{Y}_{t_k}}(F_t^{-1}(\mathbf{x})) |\nabla(F_t^{-1}(\mathbf{x}))|. \quad (\text{J.4})$$

Moreover, the forward process indicates $\mathbf{Y}_{t_k} = \mathbf{Y}_t + \mathbf{Z}_{t,t_k}$, where $\mathbf{Z}_{t,t_k} \sim N(0, t - t_k)$ is independent of \mathbf{Y}_t . Using the Jacobian transformation of probability densities, we have

$$\begin{aligned} p_{\mathbf{Y}_{t_k}}(F_t^{-1}(\mathbf{x})) &= p_{\mathbf{Y}_t + \mathbf{Z}_{t,t_k}}(F_t^{-1}(\mathbf{x})) \\ &= \int_{\mathbb{R}^d} q(T - t, \mathbf{y}) \phi_{t-t_k}(F_t^{-1}(\mathbf{x}) - \mathbf{y}) d\mathbf{y}, \end{aligned}$$

where $\phi_{\sigma^2}(\mathbf{x}) = \exp(-\|\mathbf{x}\|^2/2\sigma^2)/(2\pi\sigma^2)^{d/2}$ is the probability density function of Gaussian distribution with variance σ^2 . The last equality holds due to the formula for the sum of independent variables. We have

$$\begin{aligned} \phi_{t-t_k}(F_t^{-1}(\mathbf{x}) - \mathbf{y}) &= \phi_{t-t_k}(\mathbf{x} - \mathbf{y} + F_t^{-1}(\mathbf{x}) - \mathbf{x}) \\ &= \frac{1}{(2\pi(t - t_k))^{d/2}} e^{-\frac{\|\mathbf{x}-\mathbf{y}+F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}} \\ &= \frac{1}{(2\pi(t - t_k))^{d/2}} e^{-\frac{\|\mathbf{x}-\mathbf{y}\|^2 + \|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2 + 2\langle \mathbf{x}-\mathbf{y}, F_t^{-1}(\mathbf{x})-\mathbf{x} \rangle}{2(t-t_k)}} \\ &\geq \frac{1}{(2\pi(t - t_k))^{d/2}} e^{-\frac{(1+2/d)\|\mathbf{x}-\mathbf{y}\|_2^2 + (1+d/2)\|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}} \\ &= \frac{1}{(1+2/d)^{d/2}} \cdot \phi_{\frac{t-t_k}{1+2/d}}(\mathbf{x} - \mathbf{y}) \cdot e^{-\frac{(1+d/2)\|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}}, \quad (\text{J.5}) \end{aligned}$$

where the first inequality holds due to the Young's inequality. In the last equality, we use the definition of $\phi_{\sigma^2}(\cdot)$ with $\sigma^2 = (t - t_k)/(1 + 2/d)$. Combining (J.4) and (J.5), we have

$$\begin{aligned} p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x}) &\geq |\nabla(F_t^{-1}(\mathbf{x}))| \cdot \int_{\mathbb{R}^d} q(T - t, \mathbf{y}) \frac{1}{(1+2/d)^{d/2}} \phi_{\frac{t-t_k}{1+2/d}}(\mathbf{x} - \mathbf{y}) e^{-\frac{(1+d/2)\|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}} d\mathbf{y} \\ &\geq \frac{1}{e |\nabla F_t(F_t^{-1}(\mathbf{x}))|} p_{\mathbf{Y}_t + N(0, \frac{t-t_k}{1+2/d})}(\mathbf{x}) \cdot e^{-\frac{(1+d/2)\|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}}, \end{aligned}$$

where we use $(1 + 2/d)^{d/2} \leq e$, the fact that $\nabla(F_t^{-1}(\mathbf{x})) = [\nabla F_t(F_t^{-1}(\mathbf{x}))]^{-1}$ and the formula for the distribution of the sum of independent random variables. Then we have

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{F_t(\mathbf{Y}_{t_k})}(\mathbf{x})} \leq e |\nabla F_t(F_t^{-1}(\mathbf{x}))| \cdot e^{\frac{(1+d/2)\|F_t^{-1}(\mathbf{x})-\mathbf{x}\|_2^2}{2(t-t_k)}} \cdot \frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{\mathbf{Y}_t + N(0, \frac{t-t_k}{1+2/d})}(\mathbf{x})}$$

Since $\mathbf{Y}_t = \mathbf{X}_0 + N(0, T - t)$, by Lemma K.3, we know that

$$\frac{p_{\mathbf{Y}_t}(\mathbf{x})}{p_{\mathbf{Y}_t + N(0, \frac{t-t_k}{1+2/d})}(\mathbf{x})} \leq \left(\frac{T - t + \frac{t-t_k}{1+2/d}}{T - t} \right)^{d/2}$$

Combining the above two inequalities and we can complete the proof of Lemma I.3 □

Proof of lemma I.4. Since $\mathbf{Y}_t = \mathbf{X}_0 + \sqrt{T-t}\mathbf{Z}$, we know that:

$$\begin{aligned} \mathbb{E} e^{\lambda \|\mathbf{Y}_t\|_2^2} &\leq \mathbb{E} e^{2\lambda \|\mathbf{X}_0\|^2} \cdot \mathbb{E} e^{2\lambda(T-t)\|\mathbf{Z}\|^2} \\ &\leq e^{2\lambda R^2} \cdot e^{2\lambda(T-t)\|\mathbf{Z}\|^2} \end{aligned}$$

here (i) is because we assumed $\lambda < 1/R^2$. Moreover, we have $\lambda \leq \frac{1}{4d(T-t)}$, this implies $2\lambda(T-t) \leq \frac{1}{2}$, thus we have

$$\begin{aligned} \mathbb{E} e^{2\lambda(T-t)\|\mathbf{Z}\|_2^2} &= \int_{\mathbb{R}^d} e^{2\lambda(T-t)\|\mathbf{x}\|_2^2} (2\pi)^{-\frac{d}{2}} e^{-\frac{\|\mathbf{x}\|_2^2}{2}} d\mathbf{x} \\ &= \int_{\mathbb{R}^d} (2\pi \frac{1}{1-4\lambda(T-t)})^{-\frac{d}{2}} (1-4\lambda(T-t))^{-\frac{d}{2}} e^{-\frac{\|\mathbf{x}\|_2^2}{2\frac{1}{1-4\lambda(T-t)}}} d\mathbf{x} \\ &= (1-4\lambda(T-t))^{-\frac{d}{2}} \\ &\lesssim 1. \end{aligned}$$

This completes the proof of Lemma I.4

Proof of Lemma I.5. By tweedie's formula we know that:

$$\nabla \log q(t, \mathbf{X}_t) = -\frac{1}{t} \left(\mathbf{X}_t - \mathbb{E}_{q_0|t}(\cdot | \mathbf{X}_t) \mathbf{X}_0 \right).$$

$$\begin{aligned}\mathbb{E}\|\nabla \log q_t(\mathbf{X}_t)\|^2 &= \frac{1}{t^2} \mathbb{E}\|\mathbf{X}_t - \mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2 \\ &= \frac{1}{t^2} [\mathbb{E}\|\mathbf{X}_t\|^2 - 2\mathbb{E}\mathbf{X}_t \cdot [\mathbf{X}_0 | \mathbf{X}_t] + \mathbb{E}\|\mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2].\end{aligned}$$

Since $X_t = X_0 + \sqrt{t}Z$, we have $\mathbb{E}\|X_t\|^2 = \mathbb{E}\|X_0\|^2 + td$. Here the second order momentum is finite because our bounded support assumption. Moreover, we know that:

$$\mathbb{E} \boldsymbol{X}_t \cdot \mathbb{E}[\boldsymbol{X}_0 | \boldsymbol{X}_t] = \mathbb{E} \boldsymbol{X}_t \cdot \boldsymbol{X}_0 = \mathbb{E} (\boldsymbol{X}_0 + \sqrt{t} \boldsymbol{Z}) \cdot \boldsymbol{X}_0 = \mathbb{E} \|\boldsymbol{X}_0\|^2.$$

We next consider the trace of the covariance matrix of X_0 given X_t :

$$\text{tr} \left(\text{Cov}_{q_{0|t}(\cdot|\boldsymbol{X}_t)}(\boldsymbol{X}_0) \right) = \mathbb{E} \left[\|\boldsymbol{X}_0\|^2 | \boldsymbol{X}_t \right] - \|\mathbb{E}[\boldsymbol{X}_0 | \boldsymbol{X}_t]\|^2,$$

Thus we know that:

$$\begin{aligned} \mathbb{E} [\|\mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2] &= \mathbb{E}\|\mathbf{X}_0\|^2 - \mathbb{E} \operatorname{tr} \left(\operatorname{Cov}_{q_{0|t}(\cdot | \mathbf{X}_t)}(\mathbf{X}_0) \right) \\ &\leq \mathbb{E}\|\mathbf{X}_0\|^2. \end{aligned}$$

Putting together, we know that:

$$\mathbb{E}\|\nabla \log q(t, \mathbf{X}_t)\|^2 \leq \frac{d}{t}.$$

K AUXILIARY LEMMAS

Theorem K.1 (Reynolds Transport Theorem Leal 2007). For a function $F(t, \mathbf{x}) : \mathbb{R} \times \mathbb{R}^d \rightarrow \mathbb{R}$ that is continuously differentiable with respect to both \mathbf{x} and t , the following equality holds:

$$\frac{\partial}{\partial t} \int_{\Omega} F(t, \mathbf{x}) \, d\mathbf{x} = \int_{\Omega} \frac{\partial}{\partial t} F(t, \mathbf{x}) \, d\mathbf{x} + \int_{\partial\Omega} F(t, \mathbf{x}) \mathbf{v}(t, \mathbf{x}) \cdot \mathbf{n}(t, \mathbf{x}) \, dS,$$

where \mathbf{n} is the outward-pointing unit normal vector, \mathbf{v} is the velocity of the area element, and dS is area element.

Lemma K.2 (Lemma 6 in Benton et al. 2024). Let X_t be the OU forward process defined in (F.1). When X_0 has finite second moments, we have:

$$\mathbb{E} \|\nabla \log q(t, \mathbf{X}_t)\|^2 \leq \frac{d}{1 - \alpha_2}.$$

3024 *Proof of Lemma K.2.* This lemma is the same as Lemma 6 in Benton et al. (2024). We provide a
 3025 proof for the completeness of our paper.

3026 Denote $1 - e^{-2t}$ by σ_t . By Tweedie's formula we know that:
 3027

$$3028 \quad \nabla \log q(t, \mathbf{X}_t) = \frac{1}{\sigma_t} \left(-\mathbf{X}_t + e^{-t} \mathbb{E}_{q_{0|t}(\cdot|\mathbf{X}_t)} \mathbf{X}_0 \right).$$

3030 Taking the expectation of the square, we have

$$3031 \quad \mathbb{E} \|\nabla \log q_t(\mathbf{X}_t)\|^2 = \sigma_t^{-2} \mathbb{E} \|e^{-t} \mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t] - \mathbf{X}_t\|^2 \\ 3032 \quad = \sigma_t^{-2} [\mathbb{E} \|\mathbf{X}_t\|^2 - 2e^{-t} \mathbb{E} \mathbf{X}_t \cdot [\mathbf{X}_0 | \mathbf{X}_t] + e^{-2t} \mathbb{E} \|\mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2].$$

3034 Since $\mathbf{X}_t = e^{-t} \mathbf{X}_0 + \sqrt{\sigma_t} \mathbf{Z}$, we have $\mathbb{E} \|\mathbf{X}_t\|^2 = e^{-2t} \mathbb{E} \|\mathbf{X}_0\|^2 + \sigma_t d$. Recall that we assume that
 3035 \mathbf{X}_0 has finite second moments. Moreover, we know that:
 3036

$$3037 \quad \mathbb{E} \mathbf{X}_t \cdot \mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t] = \mathbb{E} \mathbf{X}_t \cdot \mathbf{X}_0 = \mathbb{E} (e^{-t} \mathbf{X}_0 + \sqrt{\sigma_t} \mathbf{Z}) \cdot \mathbf{X}_0 = e^{-t} \mathbb{E} \|\mathbf{X}_0\|^2.$$

3038 We next consider the trace of the covariance matrix of \mathbf{X}_0 given \mathbf{X}_t :

$$3039 \quad \text{tr} \left(\text{Cov}_{q_{0|t}(\cdot|\mathbf{X}_t)}(\mathbf{X}_0) \right) = \mathbb{E} [\|\mathbf{X}_0\|^2 | \mathbf{X}_t] - \|\mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2,$$

3041 Thus we know that:

$$3043 \quad \mathbb{E} [\|\mathbb{E}[\mathbf{X}_0 | \mathbf{X}_t]\|^2] = \mathbb{E} \|\mathbf{X}_0\|^2 - \mathbb{E} \text{tr} \left(\text{Cov}_{q_{0|t}(\cdot|\mathbf{X}_t)}(\mathbf{X}_0) \right) \\ 3044 \quad \leq \mathbb{E} \|\mathbf{X}_0\|^2.$$

3046 Putting together, we know that:

$$3048 \quad \mathbb{E} \|\nabla \log q(t, \mathbf{X}_t)\|^2 \leq \frac{d}{\sigma_t}.$$

3050 \square

3051 **Lemma K.3.** For any data distribution p_{data} and positive parameters δ and h . Let $\mathbf{X}, \mathbf{Z}_\delta, \mathbf{Z}_h$ be
 3052 three independent random variables in \mathbb{R}^d satisfying $\mathbf{X} \sim p_{\text{data}}$, $\mathbf{Z}_\delta \sim N(0, \delta)$ and $\mathbf{Z}_h \sim N(0, h)$.
 3053 Then we have:

$$3054 \quad \frac{p_{\mathbf{X}+\mathbf{Z}_\delta}(\mathbf{x})}{p_{\mathbf{X}+\mathbf{Z}_\delta+\mathbf{Z}_h}(\mathbf{x})} \leq \left(\frac{\delta+h}{\delta} \right)^{d/2}.$$

3057 where we use $p_{\mathbf{Y}}(\mathbf{x})$ to denote the probability density function of random variable \mathbf{Y} .

3058 *Proof of Lemma K.3.* We define $\phi_{\sigma^2}(\mathbf{x}) := \frac{1}{(2\pi\sigma^2)^{d/2}} e^{-\frac{\|\mathbf{x}\|^2}{2\sigma^2}}$, which is the probability density func-
 3059 tion of normal distribution $N(0, \sigma^2 \mathbf{I}_d)$.
 3060 First, we provide an upper bound of $\phi_\delta(\mathbf{x} - \mathbf{y}) / \phi_{\delta+h}(\mathbf{x} - \mathbf{y})$:

$$3063 \quad \frac{\phi_\delta(\mathbf{x} - \mathbf{y})}{\phi_{\delta+h}(\mathbf{x} - \mathbf{y})} = \left(\frac{\delta+h}{\delta} \right)^{d/2} e^{\frac{(\mathbf{x}-\mathbf{y})^2}{2(\delta+h)} - \frac{(\mathbf{x}-\mathbf{y})^2}{2\delta}} \leq \left(\frac{\delta+h}{\delta} \right)^{d/2}. \quad (\text{K.1})$$

3065 Using the independence property, we have

$$3067 \quad p_{\mathbf{X}+\mathbf{Z}_\delta}(\mathbf{x}) = \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_\delta(\mathbf{x} - \mathbf{y}) d\mathbf{y}, \\ 3068 \\ 3069 \quad p_{\mathbf{X}+\mathbf{Z}_\delta+\mathbf{Z}_h}(\mathbf{x}) = \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_{\delta+h}(\mathbf{x} - \mathbf{y}) d\mathbf{y}.$$

3071 Then we have

$$3073 \quad \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_\delta(\mathbf{x} - \mathbf{y}) d\mathbf{y} = \left(\int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_{\delta+h}(\mathbf{x} - \mathbf{y}) \frac{\phi_\delta(\mathbf{x} - \mathbf{y})}{\phi_{\delta+h}(\mathbf{x} - \mathbf{y})} d\mathbf{y} \right) \\ 3074 \\ 3075 \quad \leq \int_{\mathbb{R}^d} p_{\text{data}}(\mathbf{y}) \phi_{\delta+h}(\mathbf{x} - \mathbf{y}) d\mathbf{y} \left(\frac{\delta+h}{\delta} \right)^{d/2},$$

3077 which completes the proof of Lemma K.3. \square