# JAMUN: TRANSFERABLE MOLECULAR CONFORMA-TIONAL ENSEMBLE GENERATION WITH WALK-JUMP SAMPLING

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### Abstract

Conformational ensembles of protein structures are immensely important to understanding protein function. Current techniques for sampling ensembles are computationally inefficient, or do not transfer to systems outside their training data. We present walk-Jump Accelerated Molecular ensembles with Universal Noise (JAMUN), a step towards the goal of efficiently sampling the Boltzmann distribution of arbitrary proteins. By extending Walk-Jump Sampling to point clouds, JAMUN enables ensemble generation at orders of magnitude faster rates than traditional molecular dynamics or state-of-the-art generators. Further, JAMUN is able to predict the stable basins of small peptides that were not seen during training.

## 1 INTRODUCTION

026 Molecules are not static. They move, and these movements can be vitally important. Protein motion 027 is required for myglobin to bind oxygen and move it around the body. Miller & Phillips (2021) Drug 028 discovery on protein kinases depends on characterizing kinase conformational ensembles. Gough 029 & Kalodimos (2024) The search for druggable "cryptic pockets" requires understanding protein dynamics.Colombo (2023) However, while machine learning (ML) methods for molecular structure prediction have experienced enormous success recently, ML methods for dynamics have yet to have 031 similar impact. ML models for generating molecular ensembles are widely considered the "next frontier" (Bowman, 2024; Miller & Phillips, 2021; Zheng et al., 2023). In this work, we present 033 JAMUN (walk-Jump Accelerated Molecular ensembles with Universal Noise), a generative ML 034 model which advances this frontier by demonstrating improvements in both speed and transferability 035 over previous approaches.

While the importance of protein dynamics is well-established, it can be exceedingly difficult to 037 sufficiently sample large biomolecular systems. The most common sampling method is molecular dynamics (MD), but it is limited by the need for very short time-steps of 1-2 femtoseconds. Many important protein dynamic phenomena occur on the timescale of milliseconds. Simulating with this 040 resolution is "...equivalent to tracking the advance and retreat of the glaciers of the last Ice Age-tens 041 of thousands of years-by noting their locations each and every second." Borhani & Shaw (2012) 042 Importantly, there is nothing fundamental about this small time-step limitation; it is an artifact of 043 high-frequency motions, such as bond vibrations, that have little to no effect on protein ensembles. 044 Leimkuhler & Matthews (2015) Enhanced sampling methods have been developed in an attempt to accelerate sampling, but they often require expert user input, and, more importantly, do not address the underlying time-step problem. Other sampling methods, such as Monte Carlo-based methods, 046 exist, but have seen limited success for large biomolecular systems. Vitalis & Pappu (2009) 047

A large number of generative models have been developed to address the sampling inefficiency problems of MD using machine learning, including continuous normalizing flows, diffusion, and flow-matching.Noé et al. (2019); Arts et al. (2023); Klein et al. (2024b;a); Zheng et al. (2024); Jing et al. (2024); Kim et al. (2024) These models have been applied to a variety of MD datasets from small molecules to peptides to full proteins. However, none of these models have proven

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Figure 1: Overview of Walk-Jump Sampling.

to be transferable. They do not work well beyond their training data. (One notable exception is the recent Transferable Boltzmann Generators (Klein & Noé, 2024) model, with which we provide extensive comparison.) While the developments in this field have been immense, transferability remains a grand challenge. Without it the usefulness of any ML model is extremely limited. For a true breakthrough in this area, an ML model must be able to generate conformational ensembles for molecules that are not in its training set.

We set out to solve this problem of transferability by developing an ML model informed by the physical priors of molecular dynamics data. JAMUN is a Walk-Jump Sampler (WJS) Saremi & 094 Hyvärinen (2019) for point clouds parameterized with an SE(3)-equivariant denoiser. In WJS noise 095 is added to clean data and a denoising neural network is trained to recover the clean samples. This 096 denoiser defines the score function of the noisy manifold which we sample using Langevin dynamics (walk step) and allows us to periodically project back to the original data distribution (jump step). 098 This framework is a natural fit for MD data. In MD, unlike for natural images or other settings where generative modeling is commonly applied, we typically are interested in sampling an ensemble of representative states rather than drawing single samples. In this setting it is advantageous to 100 generate samples from trajectories that efficiently traverse the smoothed space, rather than starting 101 over from an uninformative Gaussian prior for each sample as is commonly done in diffusion and 102 flow matching. By adding partial noise, WJS simply smooths out the distribution enough to resolve 103 sampling difficulties without fully destroying the information present in the data distribution. More-104 over, the use of Langevin dynamics, the same algorithm commonly used for MD simulations, on the 105 smoothed noisy manifold lends itself to simple, physical interpretations of model behavior. 106

107 We train JAMUN on a large dataset of MD simulations of two amino acid peptides. We demonstrate that this model can generalize to a holdout set of unseen peptides. In all of these cases, generation

with JAMUN yields converged sampling of the conformational ensemble faster than MD with a standard force field. These results suggest that this transferability is a consequence of retaining the physical priors inherent in MD data. By smoothing out the underlying data distribution, JAMUN is able to produce the first transferable generative model for molecular conformational ensembles that is dramatically faster than MD simulation.

114 2 RELATED WORK

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116 The goal of building machine learning models that can generate conformational ensembles of molec-117 ular systems is not new. While a full overview of this field is beyond the scope of this work (see 118 Aranganathan et al. (2024) for a recent review), we note a few relevant previous efforts. Boltzmann 119 Generators (Noé et al., 2019) introduced the idea that a neural network could be used to transform the 120 underlying data distribution into an easier-to-sample Gaussian distribution. There have been followon efforts which used diffusion models (Arts et al., 2023), flow-matching (Klein et al., 2024b), and 121 continuous normalizing flows (Klein & Noé, 2024). The commonality in these models is the choice 122 of target distribution; they all attempt to transform the MD data distribution into a simple Gaussian. 123 This is a key difference between prior work and our model. Notably, no previous model in this area 124 except Klein & Noé (2024) has been transferable. 125

There have also been efforts to build ML models for taking longer MD time-steps (Klein et al., 2024a; Schreiner et al., 2023; Hsu et al., 2024) and for approximating conformations of large proteins (Zheng et al., 2024; Jing et al., 2024). These methods rely on hand-crafted featurizations (eg. backbone torsion angles). In practice, this has made generalization to unseen molecules challenging for these models as well.

The above models are often classified as Boltzmann Generators or Boltzmann Emulators. Models
in the former class are guaranteed to draw unbiased samples from the Boltzmann distribution, while
models in the latter class do not have this guarantee. Strictly speaking, JAMUN is a Boltzmann
Emulator, although in practical terms, as our results show, the difference is minimal.

JAMUN is a Walk-Jump Sampling method which uses an SE(3)-equivariant neural network for denoising. WJS is built on the seminal work of Neural Empirical Bayes (Saremi & Hyvärinen, 2019), and has been used in voxelized molecule generation (Pinheiro et al., 2024b;a) and protein sequence generation (Frey et al., 2023). Our work is the first to our knowledge to apply WJS to point clouds. The equivariant denoising network we use is built with the e3nn library (Geiger & Smidt, 2022). For a survey of equivariant models for 3D atomic systems, see Duval et al. (2023).

- 3 Methods
- 144 3.1 WALK-JUMP SAMPLING

145146JAMUN operates by performing Walk-Jump sampling on molecular systems represented as 3D point147clouds. A conceptual overview of the process is illustrated in Figure 1. Given an initial sample  $x^{(0)}$ 148from the clean data distribution  $p_X$  where  $x^{(0)} \in \mathbb{R}^{N \times 3}$  represents the 3D coordinates of each of149the N atoms, Walk-Jump performs the following steps:

- 1. Construct initial sample  $y^{(0)}$  from the noisy data distribution  $p_Y$  by adding noise with magnitude  $\sigma$  drawn from the normal distribution  $\mathcal{N}$ :
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- $y^{(0)} = x^{(0)} + \sigma \eta^{(0)} \text{ where } \eta^{(0)} \sim \mathcal{N}(0, \mathbb{I}_{N \times 3})$ (1)
- 2. Walk to obtain samples  $y^{(1)}, \ldots, y^{(N)}$  from  $p_Y$  using Langevin dynamics which conists of numerically solving the following Stochastic Differential Equation (SDE):

$$dy = v_y dt \tag{2}$$

$$dv_y = \nabla_y \log p_Y(y)dt - \gamma v_y dt + M^{-\frac{1}{2}}\sqrt{2}dB_t$$
(3)

159 where  $v_y$  represents the particle velocity,  $\nabla_y \log p_Y(y)$  is the gradient of the log of the 160 probability density function (the learned score function),  $\gamma$  is friction, M is the mass, and 161  $B_t$  is the standard Wiener process in  $N \times 3$ -dimensions:  $B_t \sim \mathcal{N}(0, t \mathbb{I}_{N \times 3})$ . In practice, we employ the BAOAB solver (Appendix G) to integrate Equation 2 numerically.





(4)

(5)

(6)

Figure 4: Denoising of  $y^{(t)}$  according to Equation 6 gives us new samples  $\hat{x}^{(t)}$ .

Importantly, the score function  $\nabla_y \log p_Y$  shows up in both the walk and jump steps.

#### 3.2 LEARNING TO DENOISE

In order to run Walk-Jump Sampling as outlined above, we must train a parameterized denoising model which takes in noisy samples y and outputs clean samples  $\hat{x}$ . We have the choice of modelling either the score  $\nabla_y \log p_Y$  or the denoiser  $\hat{x}$  as they are equivalent by Equation 6. Here, we follow the recommendations of Karras et al. (2022; 2024), originally developed for diffusion models in images to model the denoiser as a neural network  $\hat{x}_{\theta}(y, \sigma) \approx \hat{x}(y)$  parameterized by model parameters  $\theta$ . We appropriately modify their construction for the point cloud context as detailed in Section 3.3. Note that while the normalization is applicable for any noise level, we only need to learn a model at a **single, fixed noise level**  $\sigma$ . This is unlike training diffusion and flow-matching models where a large range of noise levels are required for sampling.

As described in Algorithm 1, we use an SE(3)-equivariant model to parametrize the denoiser 224 Thomas et al. (2018). This is in contrast to existing methods (Hoogeboom et al., 2022; Klein & Noé, 225 2024; Klein et al., 2024a) that utilize the E(3)-equivariant EGNN model (Satorras et al., 2022). As rightly pointed out by Dumitrescu et al. (2024), E(3) models are equivariant under parity, which 226 means that are forced to transform mirrored structures identically. When we experimented with 227 such architectures, we found symmetric Ramachandran plots which arise from the unnecessary par-228 ity constraint of the denoising network. For this reason, Klein & Noé (2024); Klein et al. (2024a) 229 use a 'chirality checker' to post-hoc fix the generated structures from their model; for JAMUN, such 230 post-processing is unnecessary because our model can distinguish between chiral structures. 231

Training the denoiser  $\hat{x}_{\theta}$  consists of solving the following optimization problem:

$$\theta^* = \arg\min_{\theta} \mathbb{E}_{\substack{X \sim p_X, \eta \sim \mathcal{N}(0, \mathbb{I}_{N \times 3}) \\ Y = X + \sigma\eta}} \|\hat{x}_{\theta}(Y, \sigma) - X\|^2 \tag{7}$$

to obtain  $\theta^*$ , the optimal model parameters. We approximate the expectation in Equation 7 by sampling  $X \sim p_X$  and  $\eta \sim \mathcal{N}(0, \mathbb{I}_{N \times 3})$ , as is standard in the empirical risk minimzation (ERM) (Vapnik, 1991) setting. We minimize the loss as a function of model parameters  $\theta$  using the firstorder optimizer Adam (Kingma & Ba, 2017) in PyTorch 2.0 (Ansel et al., 2024; Falcon & The PyTorch Lightning team, 2019).

241 Algorithm 1 Operations of our SE(3)-Equivariant Denoiser  $F_{\theta}$ 242 **Require:** Sample y, Noise Level  $\sigma$ , Message Passing Iterations T, Cutoff  $d_{\text{max}}$ , Spherical Harmonic 243 Degree  $\ell$ , Tensor Product  $\otimes$ 244 Compute neighbor lists for each atom in y: 245 246  $(u,v) \in E \iff ||y_u - y_v|| \le d_{\max}$ 247 for  $v \in V$  do:  $h_v^{(0)} \leftarrow \text{InitialAtomEmbedding}(v)$ 248 249 for t = 1, 2, ..., T do: 250 for  $v \in V$  do: 251  $\tilde{h}_v^{(\widetilde{t})} \leftarrow \frac{1}{|\mathcal{N}(v)|} \sum_{u \in \mathcal{N}(v)} \mathrm{MLP}(\|\mathbf{y}_u - \mathbf{y}_v\|) \times \mathrm{Linear}(h_u^{(t-1)} \otimes Y_\ell(\mathbf{r}_u - \mathbf{r}_v))$  $h_v^{(t)} \leftarrow \text{GATE}(h_v^{(t)})$ 253  $h_v^{(t)} \leftarrow \text{CONCATENATE}([h_v^{(t-1)}, h_v^{(t)}])$ 254  $h_v^{(t)} \leftarrow \text{LINEAR}(h_v^{(t)})$ 256  $h_v^{(t)} \leftarrow \text{NOISECONDITIONALSCALING}(h_v^{(t)}, \sigma)$ 257 for  $v \in V$  do: 258  $f_v \leftarrow \text{Linear}(h_v^{(T)})$ 259

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return  $\{f_v\}_{v \in V}$ 

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Similar to Klein & Noé (2024), the initial embedding of each atom (represented by INI-TIALATOMEMBEDDING in Algorithm 1) is given by embedding its atomic number, the atom type as provided by the PDB, the type of its associated residue, and the sequence index of the associated residue.

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#### 3.3 PARAMETRIZATION OF THE DENOISER NETWORK

We describe the parametrization of the denoiser network  $\hat{x}_{\theta}(y, \sigma)$  which will approximate  $\hat{x}(y)$ . While we fix a noise level  $\sigma$  here, we describe the general construction for an arbitrary noise level  $\sigma$ , which could be useful for training E(3) and SE(3) equivariant diffusion and flow-matching models as well.

We adapt the analysis and choices of Karras et al. (2022; 2024) for the point-cloud setting:

 $\hat{x}_{\theta}(y,\sigma) = c_{\text{skip}}(\sigma)y + c_{\text{out}}(\sigma)F_{\theta}(c_{\text{in}}(\sigma)y, c_{\text{noise}}(\sigma))$ 

where  $F_{\theta}$  is the SE(3)-equivariant graph neural network (GNN) model, parameterized by  $\theta$ .  $c_{skip}(\sigma), c_{out}(\sigma), c_{in}(\sigma), c_{noise}(\sigma)$  are fixed functions from  $\mathbb{R}^+$  to  $\mathbb{R}$ , chosen to normalize the effective inputs and outputs to  $F_{\theta}$ . Further, these coefficients encourage re-use of the input y at low noise levels, but the opposite at high noise levels. Next, we describe the exact form of these functions.

#### 3.4 NORMALIZATION

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As the noise level  $\sigma$  is increased,  $y = x + \sigma \eta$  where  $\eta \sim \mathcal{N}(0, \mathbb{I}_{D \times N})$  expands in space. Let  $\tilde{y}$  represent the 'normalized' input y, as seen by the network  $F_{\theta}$ :

$$\tilde{y} = c_{\rm in}(\sigma)y$$
 (9)

(8)

To control the expansion of y,  $c_{in}(\sigma)$  is chosen such that the following property holds:

$$\mathbb{E}_{\substack{(i,j) \sim \text{Uniform}(E) \\ \eta \sim \mathcal{N}(0, \mathbb{I}_{D \times N})}} \|\tilde{y}_i - \tilde{y}_j\|^2] = 1 \text{ at all noise levels } \sigma.$$
(10)

Note that this is distinct from the normalization chosen by (Karras et al., 2022; 2024), which normalizes ||y|| directly. The intuition behind this normalization is that the GNN model  $F_{\theta}$  does not operate on atom positions y directly, but instead uses the relative vectors  $y_i - y_j$  to account for translation invariance, and controlling this object directly ensures that the topology of the graph does not change with varying noise level  $\sigma$ .

To achieve this, we compute:

$$c_{\rm in}(\sigma) = \frac{1}{\sqrt{C + 2D\sigma^2}} \tag{11}$$

Where *D* is the number of particles and  $C = \mathbb{E}_{(i,j) \sim \text{Uniform}(E)} ||x_i - x_j||^2$  can be easily estimated from the true data distribution. The full derivation can be found in Appendix C.

As the input is now appropriately normalized, the target output of the network  $F_{\theta}$  should also be appropriately normalized. A full derivation, found in Appendix D, leads to:

$$c_{\rm skip}(\sigma) = \frac{C}{C + 2D\sigma^2} \tag{12}$$

$$c_{\rm out}(\sigma) = \sqrt{\frac{C \cdot 2D\sigma^2}{C + 2D\sigma^2}}$$
(13)

$$c_{\text{noise}}(\sigma) = \log_{10}\sigma \tag{14}$$

The noise normalization is a scaled version of the recommendation of  $\frac{1}{4} \ln \sigma$  for images in Karras et al. (2022; 2024).

# 312313 3.5 ROTATIONAL ALIGNMENT

As described in Algorithm 2, we use the Kabsch-Umeyama algorithm (Kabsch, 1976; Umeyama, 1991) to rotationally align y to x before calling the denoiser.

Note that both y and x are mean-centered to respect translational equivariance:

$$\sum_{i=1}^{N} y_i = \vec{0} \in \mathbb{R}^3 \tag{15}$$

$$\sum_{i=1}^{N} x_i = \vec{0} \in \mathbb{R}^3$$
(16)

so there is no net translation.

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<b>Require:</b> Noisy Sample $y \in \mathbb{R}^{N \times 3}$ , True Sample $x \in \mathbb{R}^{N \times 3}$ .	
$\hat{H} \leftarrow x^T y$	$\triangleright H \in \mathbb{R}^{3 \times 3}$
$U, S, V^T \leftarrow \text{SVD}(H)$	$\triangleright U, V \in \mathbb{R}^{3 \times 3}$
$\mathbf{R}^* \leftarrow U \operatorname{diag}[1, 1, \operatorname{det}(U) \operatorname{det}(V)] V^T$ return $y(\mathbf{R}^*)^T$	

#### 4 DATASET

We demonstrate our method on two datasets: (i) the uncapped amino acids used in Timewarp (Klein et al., 2024a), specifically the 2AA-huge dataset consisting of 380 diamines split into 200 train, 80 validation and 100 test diamines, and (ii) capped amino acids simulated in water following the same splits used in (i). The molecules in (i) have a single peptide bond between amino acids of varying identities. This results in a distribution well represented by two dihedral angles, the  $\phi$  angle of the second residue and the  $\psi$  angle of the first residue. The termini are zwitterionic amino and carboxyl groups. These are not ideal analogues of amino acids in proteins due to local charge interactions as well as lack of steric effects. However, for consistency in benchmarking, we run all experiments on these to compare against Transferable Boltzman Generators (Klein & Noé, 2024).



Figure 5: A side-by-side comparison of uncapped (left) compared to capped (right) ALA-CYS. The acetyl (ACE) and N-methyl (NME) capping groups provide steric hindrance and prevent local charge interactions on the N-terminal and C-terminal ends respectively.

As is common in molecular dynamics simulations of very small peptides, we also generate a similar dataset with ACE (acetyl) and NME (N-methyl amide) caps. These are essentially peptide bonds with the first and last residue, bonded to methyl groups. These peptide bonds remove the need for the zwitterion, while the methyl group provides some steric interactions. This results in a distribution which requires at least 4 dihedrals to be well-represented: the  $\phi$  and  $\psi$  angles of both residues. We find this is a significantly richer set of distributions. We ensure that our unbiased molecular dynamics runs are converged or representative by comparing against biased molecular dynamics runs using Non-equilibrium Umbrella Sampling (NEUS)Dinner et al. (2018); Vani et al. (2022), a trajectory stratification based enhanced sampling algorithm. 

## 5 Results

We find that JAMUN samples conformational ensembles of short peptides faster than conventional MD while remaining reasonably faithful to thermodynamics. It does this despite being quite small by ML model standards (8.2M parameters). We note it requires no a priori knowledge or hand-crafted featurization for transferable exploration–for instance, it generates diversity in  $\phi$  and  $\psi$  dihedrals without explicit inputs regarding these variables. We also show a remarkable speedup relative to transferable Boltzmann generators. However, unlike the transferable Boltzmann generator, we only have access to the score of the distribution, and thus cannot precisely reweight it, making this a Boltzmann emulator.



Figure 6: -log probabilities of outputs from four example diamines from the test set. These, along
with Figure 9, are chosen to represent the diversity of states as well as subtle perturbations from
single-point amino acid changes in the dataset. Each pair of Ramachandran plots show the distribution for a single diamine, with each row corresponding to diamine identity. The first column shows
the fully converged molecular dynamics distribution. The second shows sixty four thousand samples
from JAMUN. The third column shows a distribution obtained from an MD trajectory run for the
exact same amount of time as the JAMUN sampler.

410 In Figure 6 and Figure 9 we show, for eight test peptides in the challenging capped dataset, distri-411 butions generated using both molecular dynamics and JAMUN. As the molecular dynamics serves 412 as our ground truth, we show the full dataset. However, for demonstrative purposes, we also show 413 the distribution when it has been run for the same amount of GPU time as JAMUN requires to sam-414 ple 640000 points. This number was chosen to reflect reasonable sampling for validation diamines. 415 This can be variable, for the specific 8 diamines we choose here, the sampling time with JAMUN 416 is: 1) ASP-TRP: 38 minutes, 2) GLU-THR: 27 minutes, 3) PHE-ALA: 28 minutes, 4) ASN-GLU: 417 29 minutes, 5) CYS-TRP: 34 minutes, 6) GLY-ASN: 22 minutes, 7) HIS-PRO: 30 minutes, and 8) ILE-GLY: 22 minutes. In terms of MD simulation time, this is between 100ps and 300ps. 418

These results clearly show that JAMUN is sampling the conformational ensembles of unseen peptides with high accuracy. The vast majority of low-energy basins in the Ramachandran plot are
captured by the model. Moreover, JAMUN is sampling these states much faster than conventional
MD. Comparing the "JAMUN" and "Benchmark Molecular Dynamics" columns of Figure 6 and
Figure 9 shows that MD lags dramatically behind. At the point where JAMUN has sampled the
entire distribution, MD is often still stuck sampling a single basin.

While the capped dataset represents higher diversity and fidelity to true peptide thermodynamics, we
also wish to compare against the transferable Boltzmann generator, as it has already shown success
for short peptides. To this end, we also train our model on the timewarp dataset, although please note
that our hyperparameters are optimized for the capped diamines. Nevertheless, we show in Figure 7
a significant improvement in sampling. For the two machine learning methods, we simulate using
an equal amount of compute-time, including post-processing steps such as chirality checks. It is
clear, by visual inspection, that JAMUN samples all states, where transferable Boltzmann Generators



Figure 7: -log probabilities of outputs from two example diamines from the uncapped test set. These are chosen to show two examples of different distributions from this dataset which has limited conformational diversity. Each Ramachandran plot shows the distribution for a single diamine, with each row corresponding to diamine identity. The first column shows the full distribution from the timewarp dataset. The second shows sixty four thousand samples from JAMUN. The third column shows a distribution obtained from a transferable Boltzman generator run for the exact same amount of time as the JAMUN sampler. The fourth shows results from the transferable Boltzman generator run for five thousand samples (run for roughly 10 times as long as JAMUN). 



Figure 8: Shannon Jensen divergences plotted along time. (left) A comparison between JAMUN and molecular Dynamics where the x-axis is a trajectory progress coordinate such that all JAMUN runs are for 640,000 samples while molecular Dynamics runs are truncated to be the same GPU time. (right) Shannon-Jensen divergence of a long fully converged molecular dynamics trajectory. 

As a convergence metric, we look at Shannon-Jensen divergence in Figure 8. This allows us to compare the rate at which JAMUN and MD converge to the "correct" distribution. We calculate the Shannon-Jensen divergence for molecular dynamics and JAMUN runs of equal GPU time (using the data generated for Figure 6 and Figure 9). We also analyze fully converged molecular dynamics runs for reference. It appears that while JAMUN converges quicker than MD, and samples relevant states faster, it does not ever converge to the same accuracy. It is likely that part of the discrepency is from JAMUN's oversampling of rare and transition regions. In particular, we see from the Ramachandran plots that occasionally the standard transition paths are not the most frequent and instead spurious transition paths are sampled along. JAMUN is not trained on kinetics and cannot faithfully reproduce them. This is a well-understood phenomena in stochastic processes in a rough to smooth
surface.Zwanzig (1988) It is also evident from the smooth nature of JAMUN's Shannon-Jensen time
series, as compared to the step-wise behavior of MD that JAMUN has a more natural mixing of
states without kinetic traps. However, much of the spurious sampling can be avoided with smaller
noise values at the cost of sampling efficiency. As such, this Langevin dynamics on a noised surface
provides an interesting avenue of research from a stochastic processes point of view.

492 From a biological and drug-design standpoint, we are relatively uninterested in kinetics or even the 493 exact distribution. However, we are extremely interested in the sampling of diverse metastable states 494 with some sense of their relative stabilities. To this end, JAMUN does an impressive job of sampling 495 metastable states particular to specific diamines with thermodynamic fidelity. We demonstrate this with the use of Markov State Models trained on converged MD data in Appendix B. The figures 496 demonstrate that across the test set of capped peptides JAMUN (i) captures the precise shape of states 497 extremely well, and (ii) samples nuanced states that are not obvious by eye in a Ramachandran plot 498 with high accuracy. 499

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## 6 CONCLUSION

We present JAMUN, a Walk-Jump Sampling model for generating ensembles of molecular conformations. The model is trained on molecular dynamics data from two amino acid peptides and
is transferable to peptides outside of its training set. It produces accurate conformational ensembles faster than MD or previous ML approaches. This represents an important first step toward the
ultimate goal of a transferable generative model for protein conformational ensembles.

The model has some limitations that motivate future work. While it is highly transferable in the space of two amino acid peptides, we have not tested the model's ability to transfer to larger proteins. We leave exploring this important direction to future work. Additionally, while the current SE(3)equivariant denoiser architecture works well, further development of the denoising network could speed up sampling. Alternative "jump" methods, such as Walk-Diffuse, could also serve to sharpen generation. Lastly, a promising direction that has not yet been explored is the application of classical enhanced sampling methods, such as metadynamics, for traversing the noisy space.

JAMUN is the first model to use Walk-Jump Sampling on point clouds, and the first to apply Walk-Jump Sampling to molecular dynamics data. This paradigm gives the model a clear physics interpretation. By adding noise, JAMUN is able to simulate on a smoother manifold than ordinary MD.
Sampling on this smooth surface enables the model to generate accurate molecular conformational
ensembles faster than both MD and previous ML models.

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A CODE AVAILABILITY

- We will be releasing all the code and data necessary to reproduce this paper.
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## **B** DISTRIBUTION ANALYSES

We demonstrate the quantitative gain in sampling through metrics like the Shannon-Jensen diver-682 gence and the clear qualitative gain in sampling diversity. Here we show this visually by employ-683 ing Markov state models (MSMs) Pande et al. (2010) to characterize metastability. MSMs are the 684 most commonly used clustering algorithm for times series in molecular dynamics. They function 685 by first projecting on a low-dimensional manifold, most commonly the Time-lagged Independent 686 Component Analysis (TICA), finely clustering the samples and then solving for the eigenvalues of a 687 transition probability matrix. Here we use the PYEMMA2 Scherer et al. (2015) implementation. We 688 use a lag time of ten steps, five tics, 200 k-means clusters to obtain 10 macrostates for the unbiased 689 well-sampled molecular dynamics trajectories. We perform an implied timescales analysis to ensure 690 that these states are a reasonable representation. We can then use this representation to better under-691 stand the features of our space and qualitatively evaluate JAMUN. In figures 10-16, we plot PMFs 692 (potential of mean force) in grayscale for JAMUN's samples with the Markov state model histogram superimposed in color. Each pair of Ramachandran plots correspond to a single metastable state. 693

In particular, note that oblong metastable basins corresponding to the  $-\pi$  region in  $\phi$ , small metastability in the  $(-\pi, \pi)$  area and even the specific shapes of commonly occuring basins can be signatures particular to specific diamine pairs and are well-emulated by JAMUN.

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C INPUT NORMALIZATION

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Fix an  $(i, j) \in E$  from Equation 10. As  $\eta_i, \eta_j \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \mathbb{I}_D)$ , we have  $\eta_i - \eta_j \sim \mathcal{N}(0, 2\mathbb{I}_D)$  from the closure of the multivariate Gaussian under linear combinations. Thus, for each component



Figure 9: -log probabilities of outputs from four example diamines from the test set. These, along with 6, are chosen to represent the diversity of states as well as subtle perturbations from single-point amino acid changes in the dataset. Each pair of Ramachandran plots show the distribution for a single diamine, with each row corresponding to diamine identity. The first column shows the fully converged molecular dynamics distribution. The second shows sixty four thousand samples from JAMUN. The third column shows a distribution obtained from an MD trajectory run for the exact same amount of time as the JAMUN sampler.



Figure 10: Markov states superimposed on sampled PMFs for CYS-TRP. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.



Figure 11: Markov states superimposed on sampled PMFs for ASP-THR. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.



Figure 12: Markov states superimposed on sampled PMFs for GLU-THR. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.

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Figure 13: Markov states superimposed on sampled PMFs for PHE-ALA. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.



Figure 14: Markov states superimposed on sampled PMFs for GLY-ASN. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.

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Figure 15: Markov states superimposed on sampled PMFs for HIS-PRO. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.

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Figure 16: Markov states superimposed on sampled PMFs for ILE-GLY. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.

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Figure 17: Markov states superimposed on sampled PMFs for ASN-GLU. This figure shows the results of running JAMUN with Markov state model negative log histograms superimposed on them. The MSMs are trained on fully converged MD data and represent metastable states. Each pair of Ramachandran plots correspond to a single metastable basin. Note that the gray-scale represents JAMUN samples while the colored is from molecular dynamics.

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1188  $d \in \{1, \ldots D\}$ , we have:  $(\eta_i - \eta_j)_{(d)} \sim \mathcal{N}(0, 2)$  and hence: 1189  $\mathbb{E}_{\eta \sim \mathcal{N}(0, \mathbb{I}_{D \times N})}[(x_i - x_j)^T (\eta_i - \eta_j)] = \sum_{l=1}^{D} (x_i - x_j)_{(d)} \mathbb{E}[(\eta_i - \eta_j)_{(d)}] = 0$ 1190 (17)1191 1192  $\mathbb{E}_{\eta \sim \mathcal{N}(0, \mathbb{I}_{D \times N})}[\|\eta_i - \eta_j\|^2] = \sum_{l=1}^{D} \mathbb{E}[(\eta_i - \eta_j)^2_{(d)}] = 2D$ 1193 (18)1194 1195 We can now compute: 1196  $\mathbb{E}_{z}[\|\tilde{y}_{i} - \tilde{y}_{i}\|^{2}]$ 1197 1198  $= c_{\rm in}(\sigma)^2 \mathbb{E}_z[\|y_i - y_i\|^2]$ 1199  $= c_{in}(\sigma)^2 \mathbb{E}_z[||x_i - x_i + \sigma(\eta_i - \eta_i)||^2]$ 1201  $= c_{\rm in}(\sigma)^2 \left( \|x_i - x_j\|^2 + 2\sigma \mathbb{E}_z[(x_i - x_j)^T (\eta_i - \eta_j)] + \sigma^2 \mathbb{E}_z[\|\eta_i - \eta_j\|^2] \right)$ 1202  $= c_{\rm in}(\sigma)^2 \left( \|x_i - x_j\|^2 + \sigma^2 \mathbb{E}_z[\|\eta_i - \eta_j\|^2] \right)$ 1203  $= c_{\rm in}(\sigma)^2 \left( \|x_i - x_j\|^2 + 2D\sigma^2 \right).$ 1205 (19)

1207 Now, taking the expectation over all  $(i, j) \in E$  uniformly:

$$\mathbb{E}_{\substack{(i,j)\sim\text{Uniform}(E)\\\eta\sim\mathcal{N}(0,\mathbb{I}_{D\times N})}} [\|\tilde{y}_i - \tilde{y}_j\|^2] = \mathbb{E}_{\substack{(i,j)\sim\text{Uniform}(E)\\ = c_{\text{in}}(\sigma)^2}} [\mathbb{E}_z[\|\tilde{y}_i - \tilde{y}_j\|^2]]$$
$$= c_{\text{in}}(\sigma)^2 \left(\mathbb{E}_{\substack{(i,j)\sim\text{Uniform}(E)\\ = m}} \|x_i - x_j\|^2 + 2D\sigma^2\right)$$
(20)

1212 Let  $C = \mathbb{E}_{(i,j) \sim \text{Uniform}(E)} ||x_i - x_j||^2$ , which we estimate from the true data distribution. Then, 1213 from Equation 20 and our intended normalization given by Equation 10:

C

 $= \mathbb{E}_{\sigma \sim p_{\sigma}} \mathbb{E}_{x \sim p_{\tau}} \mathbb{E}_{\eta \sim \mathcal{N}(0,\mathbb{I})} [\lambda(\sigma) \| \hat{x}_{\theta}(x + \sigma\eta, \sigma) - x \|^{2}$ 

 $= \mathbb{E}_{\sigma \sim p_{\sigma}} \mathbb{E}_{x \sim p_{x}} \mathbb{E}_{y \sim \mathcal{N}(x, \sigma^{2}\mathbb{I})} [\lambda(\sigma) \| \hat{x}_{\theta}(y, \sigma) - x \|^{2}]$ 

$$c_{\rm in}(\sigma) = \frac{1}{\sqrt{C + 2D\sigma^2}} \tag{21}$$

### 1217 1218 D OUTPUT NORMALIZATION

 $\mathcal{L}(\hat{x}_{\theta}) = \mathbb{E}_{\sigma \sim p_{\sigma}}[\lambda(\sigma)\mathcal{L}(\hat{x}_{\theta}, \sigma)]$ 

The derivation here is identical that of (Karras et al., 2022; 2024), but with our normalization. The denoising loss at a single noise level is:

$$\mathcal{L}(\hat{x}_{\theta},\sigma) = \mathbb{E}_{x \sim p_x} \mathbb{E}_{\eta \sim \mathcal{N}(0,\mathbb{I})} [\|\hat{x}_{\theta}(x+\sigma\eta,\sigma)-x\|^2]$$
(22)

which gets weighted across a distribution  $p_{\sigma}$  of noise levels by (unnormalized) weights  $\lambda(\sigma)$ :

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$$= \mathbb{E}_{\sigma \sim p_{\sigma}} \mathbb{E}_{x \sim p_{x}} \mathbb{E}_{y \sim \mathcal{N}(x, \sigma^{2}\mathbb{I})} [\lambda(\sigma) \| c_{\text{skip}}(\sigma) y + c_{\text{out}}(\sigma) F_{\theta}(c_{\text{in}}(\sigma) y, c_{\text{noise}}(\sigma)) - x \|^{2}]$$

$$= \mathbb{E}_{\sigma \sim p_{\sigma}} \mathbb{E}_{x \sim p_{x}} \mathbb{E}_{y \sim \mathcal{N}(x, \sigma^{2}\mathbb{I})} \left[ \lambda(\sigma) c_{\text{out}}(\sigma)^{2} \left\| F_{\theta}(c_{\text{in}}(\sigma) y, c_{\text{noise}}(\sigma)) - \frac{x - c_{\text{skip}}(\sigma) y}{c_{\text{out}}(\sigma)} \right\|^{2} \right]$$

$$= \mathbb{E}_{\sigma \sim p_{\sigma}} \mathbb{E}_{x \sim p_{x}} \mathbb{E}_{y \sim \mathcal{N}(x, \sigma^{2}\mathbb{I})} \left[ \lambda(\sigma) c_{\text{out}}(\sigma)^{2} \left\| F_{\theta}(c_{\text{in}}(\sigma) y, c_{\text{noise}}(\sigma)) - F \right\|^{2} \right]$$
(23)

1235 where:

$$F = \frac{x - c_{\rm skip}(\sigma)y}{c_{\rm out}(\sigma)}$$
(24)

is the effective training target for the network  $F_{\theta}$ . We want to normalize F similarly as the network input:

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$$\mathbb{E}_{\substack{(i,j) \sim \text{Uniform}(E)\\\eta \sim \mathcal{N}(0, \mathbb{I}_{D \times N})}} [\|F_i - F_j\|^2] = 1 \text{ at all noise levels } \sigma.$$
(25)

Again, for a fixed  $(i, j) \in E$ , we have: 

 $\mathbb{E}_{z} \|F_{i} - F_{j}\|^{2} = \frac{\mathbb{E}_{z} \|(x_{i} - x_{j}) - c_{\text{skip}}(\sigma)(y_{i} - y_{j})\|^{2}}{c_{\text{out}}(\sigma)^{2}}$  $= \frac{\mathbb{E}_z \left\| (1 - c_{\text{skip}}(\sigma))(x_i - x_j) - c_{\text{skip}}(\sigma)\sigma \cdot (\eta_i - \eta_j) \right\|^2}{c_{\text{out}}(\sigma)^2}$  $= \frac{(1 - c_{\text{skip}}(\sigma))^2 \left\| x_i - x_j \right\|^2 + c_{\text{skip}}(\sigma)^2 \cdot 2D\sigma^2}{c_{\text{out}}(\sigma)^2}$ 

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$$\mathbb{E}_{(i,j)\sim \text{Uniform}(E)}[||F_i - F_j||^2] = 1$$

$$\eta \sim \mathcal{N}(0,\mathbb{I}_{D \times N})$$

$$\implies \frac{(1 - c_{\text{skip}}(\sigma))^2 \cdot C + c_{\text{skip}}(\sigma)^2 \cdot 2D\sigma^2}{c_{\text{out}}(\sigma)^2} = 1$$

$$\implies c_{\text{out}}(\sigma)^2 = (1 - c_{\text{skip}}(\sigma))^2 \cdot C + c_{\text{skip}}(\sigma)^2 \cdot 2D\sigma^2 \qquad (27)$$

where C was defined above. Now, to minimize  $c_{out}(\sigma)$  to maximize reuse and avoid amplifying network errors, as recommended by Karras et al. (2022; 2024): 

$$\frac{d}{dc_{\text{skip}}(\sigma)}c_{\text{out}}(\sigma)^{2} = 0$$

$$\implies -2(1 - c_{\text{skip}}(\sigma)) \cdot C + 2c_{\text{skip}}(\sigma) \cdot 2D\sigma^{2} = 0$$

$$\implies c_{\text{skip}}(\sigma) = \frac{C}{C + 2D\sigma^{2}}$$
(28)

Substituting into Equation 27, we get after some routine simplification:

$$c_{\rm out}(\sigma) = \sqrt{\frac{C \cdot 2D\sigma^2}{C + 2D\sigma^2}}$$
(29)

(26)

The noise normalization is chosen as  $c_{\text{noise}}(\sigma) = \log_{10} \sigma$ , a scaled version of the recommendation of  $\frac{1}{4} \ln \sigma$  for images in Karras et al. (2022; 2024). 

From Equation 23, we set  $\lambda(\sigma) = \frac{1}{c_{\text{out}}(\sigma)^2}$ . 

#### Ε THE DENOISER MINIMIZES THE EXPECTED LOSS

Here, we prove Equation 5, rewritten here for clarity: 

$$\hat{x}(\cdot) \equiv \mathbb{E}[X \mid Y = \cdot] = \underset{a:\mathbb{R}^{N\times3}\to\mathbb{R}^{N\times3}}{\operatorname{arg\,min}} \mathbb{E}_{\substack{X\sim p_X,\eta\sim\mathcal{N}(0,\mathbb{I}_{N\times3})\\Y=X+\sigma\eta}} [\|a(Y) - X\|^2]$$
(30)

First, we can decompose the loss over the domain  $\mathbb{R}^{N \times 3}$  of Y: 

$$\mathbb{E}_{X \sim p_X, \eta \sim \mathcal{N}(0, \mathbb{I}_{N \times 3})} [\|a(Y) - X\|^2] = \mathbb{E}_{X \sim p_X, Y \sim p_Y} [\|a(Y) - X\|^2]$$

$$Y = X + \sigma \eta$$
(31)

$$\int_{\mathbb{R}^{N\times3}} \int_{\mathbb{R}^{N\times3}} \|a(y) - x\|^2 p_{X,Y}(x,y) dxdy$$
(32)

$$= \int_{\mathbb{R}^{N\times3}} \underbrace{\int_{\mathbb{R}^{N\times3}} \|a(y) - x\|^2 p_{Y|X}(y \mid x) p_X(x) dx}_{l(a,y)} dy \quad (33)$$

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$$= \int_{\mathbb{R}^{N\times 3}} l(a, y) dy$$
 (34)  
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=

and hence, any minimizer  $a^*$  must minimize the local denoising loss  $l(a^*, y)$  at each point  $y \in \mathbb{R}^{N \times 3}$ . For a fixed  $y \in \mathbb{R}^{N \times 3}$ , the loss l(a, y) is convex, and hence, the global minimizer can be 

found by finding the critical points of l(a, y) as a function of a(y):  $\nabla_{a(y)}l(a,y) = 0$ (35) $\implies \nabla_{a(y)} \int_{\mathbb{R}^{N \times 3}} \left\| a(y) - x \right\|^2 p_{Y|X}(y \mid x) p_X(x) dx = 0$ (36) $\implies \int_{\mathbb{R}^{N\times 3}} 2(a^*(y) - x) p_{Y|X}(y \mid x) p_X(x) dx = 0$ (37)Rearranging:  $a^*(y) = \frac{\int_{\mathbb{R}^{N\times 3}} x p_{Y|X}(y \mid x) p_X(x) dx}{\int_{\mathbb{R}^{N\times 3}} p_{Y|X}(y \mid x) p_X(x) dx}$ (38) $= \frac{\int_{\mathbb{R}^{N\times 3}} x \, p_{Y|X}(y \mid x) p_X(x) dx}{p_Y(y)}$  $= \int_{\mathbb{R}^{N\times 3}} x \, \frac{p_{Y|X}(y \mid x) p_X(x)}{p_Y(y)} dx$ (39)(40) $= \int_{\mathbb{R}^{N\times 3}} x \, p_{X|Y}(x \mid y) dx$ (41) $= \mathbb{E}[X \mid Y = y]$ (42) $= \hat{x}(y)$ (43)by Bayes' rule. Hence, the denoiser as defined by Equation 4 is indeed the minimzer of the denoising loss:

$$\hat{x}(\cdot) \equiv \mathbb{E}[X \mid Y = \cdot] = \underset{a:\mathbb{R}^{N\times3} \to \mathbb{R}^{N\times3}}{\operatorname{argmin}} \mathbb{E}_{X \sim p_X, \eta \sim \mathcal{N}(0, \mathbb{I}_{N\times3})} [\|a(Y) - X\|^2]$$
(44)

as claimed. 

#### F **RELATING THE SCORE AND THE DENOISER**

Here, we rederive Equation 6, relating the score function  $\nabla \log p_Y$  and the denoiser  $\hat{x}$ , as first shown by Robbins (1956); Miyasawa (1960). 

Let  $X \sim p_X$  defined over  $\mathbb{R}^{N \times 3}$  and  $\eta \sim \mathcal{N}(0, \mathbb{I}_{N \times 3})$ . Let  $Y = X + \sigma \eta$ , which means: 

$$p_{Y|X}(y \mid x) = \mathcal{N}(y; x, \mathbb{I}_{N \times 3}) = \frac{1}{(2\pi\sigma^2)^{\frac{3N}{2}}} \exp\left(-\frac{\|y - x\|^2}{2\sigma^2}\right)$$
(45)

Then:

$$\mathbb{E}[X \mid Y = y] = y + \sigma^2 \nabla_y \log p_Y(y) \tag{46}$$

To prove this:

$$\nabla_y p_{Y|X}(y \mid x) = -\frac{y - x}{\sigma^2} p_{Y|X}(y \mid x) \tag{47}$$

$$\implies (x - y)p_{Y|X}(y \mid x) = \sigma^2 \nabla_y p_{Y|X}(y \mid x)$$
(48)

$$\implies \int_{\mathbb{R}^{N\times 3}} (x-y) p_{Y|X}(y\mid x) \ p_X(x) dx = \int_{\mathbb{R}^{N\times 3}} \sigma^2 \nabla_y p_{Y|X}(y\mid x) \ p_X(x) dx \qquad (49)$$

By Bayes' rule: 

$$p_{Y|X}(y \mid x)p_X(x) = p_{X,Y}(x,y) = p_{X|Y}(x|y)p_Y(y)$$
(50)

and, by definition of the marginals: 

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$$\int_{\mathbb{R}^{N\times 3}} p_{X,Y}(x,y) dx = p_Y(y)$$
(51)

For the left-hand side, we have:  $\int_{\mathbb{R}^{N\times3}} (x-y)p_{Y|X}(y \mid x)p_X(x)dx = \int_{\mathbb{R}^{N\times3}} (x-y)p_{X,Y}(x,y)dx \qquad (52)$   $= \int_{\mathbb{R}^{N\times3}} xp_{X,Y}(x,y)dx - \int_{\mathbb{R}^{N\times3}} yp_{X,Y}(x,y)dx \qquad (53)$   $= p_Y(y) \left( \int_{\mathbb{R}^{N\times3}} xp_{X|Y}(x \mid y)dx - y \int_{\mathbb{R}^{N\times3}} p_{X|Y}(x \mid y)dx \right) \qquad (54)$   $= p_Y(y) \left( \mathbb{E}[X \mid Y = y] - y \right) \qquad (55)$ For the right-hand side, we have:

$$\sigma^2 \int_{\mathbb{R}^{N\times3}} \nabla_y p_{Y|X}(y \mid x) p_X(x) dx = \sigma^2 \nabla_y \int_{\mathbb{R}^{N\times3}} p_{Y|X}(y \mid x) p_X(x) dx \tag{56}$$

$$=\sigma^2 \nabla_y \int_{\mathbb{R}^{N\times 3}} p_{X,Y}(x,y) dx \tag{57}$$

$$=\sigma^2 \nabla_y p_Y(y) \tag{58}$$

1368 Thus,

$$p_Y(y)\left(\mathbb{E}[X \mid Y=y] - y\right) = \sigma^2 \nabla_y p_Y(y) \tag{59}$$

$$\Rightarrow \mathbb{E}[X \mid Y = y] = y + \sigma^2 \frac{\nabla_y p_Y(y)}{p_Y(y)}$$
(60)

$$= y + \sigma^2 \nabla_y \log p_Y(y) \tag{61}$$

1374 as claimed.

# G NUMERICAL SOLVERS FOR LANGEVIN DYNAMICS

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As mentioned in Section 3.1, solving the Stochastic Differential Equation corresponding to Langevin dynamics is often performed numerically. In particular, BAOAB (Leimkuhler & Matthews, 2012; 2015; Sachs et al., 2017) refers to a 'splitting method' that solves the Langevin dynamics SDE by splitting it into three different components labelled by  $\mathcal{A}$ ,  $\mathcal{B}$  and  $\mathcal{O}$  below:

$$dy = \underbrace{v_y dt}_{\mathcal{A}} \tag{62}$$

$$dv_y = \underbrace{M^{-1} \nabla_y \log p_Y(y) dt}_{\mathcal{B}} - \underbrace{\gamma v_y dt + \sqrt{2\gamma} M^{-\frac{1}{2}} dB_t}_{\mathcal{O}}$$
(63)

1388 where both  $y, v_y \in \mathbb{R}^d$ . This leads to the following update operators:

$$\mathcal{A}_{\Delta t} \begin{bmatrix} y \\ v_y \end{bmatrix} = \begin{bmatrix} y + v_y \Delta t \\ v_y \end{bmatrix}$$
(64)

$$\mathcal{B}_{\Delta t} \begin{bmatrix} y \\ v_y \end{bmatrix} = \begin{bmatrix} y \\ v_y + M^{-1} \nabla_y \log p_Y(y) \Delta t \end{bmatrix}$$
(65)

$$\mathcal{O}_{\Delta t} \begin{bmatrix} y\\ v_y \end{bmatrix} = \begin{bmatrix} y\\ e^{-\gamma \Delta t}v_y + M^{-\frac{1}{2}}\sqrt{1 - e^{-2\gamma \Delta t}}B \end{bmatrix}$$
(66)

1396 where  $B \sim \mathcal{N}(0, \mathbb{I}_d)$  is resampled every iteration. As highlighted by Kieninger & Keller (2022), the 1397  $\mathcal{A}$  and  $\mathcal{B}$  updates are obtained by simply discretizing the updates highlighted in Equation 62 by the 1398 Euler method. The  $\mathcal{O}$  update refers to a explicit solution of the Ornstein-Uhlenbeck process, which 1399 we rederive for completeness in Appendix H.

Finally, the iterates of the BAOAB algorithm are given by a composition of these update steps matching the name of the method:

 $\begin{bmatrix} y^{(t+1)} \\ v_y^{(t+1)} \end{bmatrix} = \mathcal{B}_{\frac{\Delta t}{2}} \mathcal{A}_{\frac{\Delta t}{2}} \mathcal{O}_{\Delta t} \mathcal{A}_{\frac{\Delta t}{2}} \mathcal{B}_{\frac{\Delta t}{2}} \begin{bmatrix} y^{(t)} \\ v_y^{(t)} \end{bmatrix}$ (67)

# 1404 H THE ORNSTEIN-UHLENBECK PROCESS

For completeness, we discuss the distributional solution of the Ornstein-Uhlenbeck process, taken directly from the excellent Leimkuhler & Matthews (2015). The Ornstein-Uhlenbeck Process corresponds to the following Stochastic Differential Equation (SDE):

$$dv_y = -\gamma v_y dt + \sqrt{2\gamma} M^{-\frac{1}{2}} dB_t \tag{68}$$

Multiplying both sides by the integrating factor  $e^{\gamma t}$ :

\_

$$e^{\gamma t} dv_y = -\gamma e^{\gamma t} (v_y dt + e^{\gamma t} \sqrt{2\gamma} M^{-\frac{1}{2}} dB_t$$
(69)

$$\Rightarrow e^{\gamma t} (dv_y + \gamma v_y dt) = e^{\gamma t} \sqrt{2\gamma} M^{-\frac{1}{2}} dB_t$$
(70)

1415 1416 and identifying:

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$$e^{\gamma t}(dv_y + \gamma v_y dt) = d(e^{\gamma t}v_y) \tag{71}$$

We get after integrating from  $t_1$  to  $t_2$ , two adjacent time steps of our integration grid:

$$d(e^{\gamma t}v_y) = e^{\gamma t}\sqrt{2\gamma}M^{-\frac{1}{2}}dB_t$$
(72)

$$\implies \int_{t_1}^{t_2} d(e^{\gamma t} v_y) = \int_{t_1}^{t_2} e^{\gamma t} \sqrt{2\gamma} M^{-\frac{1}{2}} dB_t \tag{73}$$

$$\implies e^{\gamma t_2} v_y(t_2) - e^{\gamma t_1} v_y(t_1) = \sqrt{2\gamma} M^{-\frac{1}{2}} \int_{t_1}^{t_2} e^{\gamma t} dB_t$$
(74)

1426 1427 1428 Now, for a Wiener process  $B_t$ , if g(t) is a deterministic function,  $\int_{t_1}^{t_2} g(t) dB_t$  is distributed as 1428  $\mathcal{N}\left(0, \int_{t_1}^{t_2} g(t)^2 dt\right)$  by Itô's integral. Thus, applying this result to  $g(t) = e^{\gamma t}$ , we get:

$$e^{\gamma t_2} v_y(t_2) - e^{\gamma t_1} v_y(t_1) = \sqrt{2\gamma} M^{-\frac{1}{2}} \mathcal{N}\left(0, \frac{e^{2\gamma t_2} - e^{2\gamma t_1}}{2\gamma}\right)$$
(75)

$$\implies v_y(t_2) = e^{-\gamma(t_2 - t_1)} v_y(t_1) + \sqrt{2\gamma} M^{-\frac{1}{2}} e^{-\gamma t_2} \mathcal{N}\left(0, \frac{e^{2\gamma t_2} - e^{2\gamma t_1}}{2\gamma}\right)$$
(76)

$$=e^{-\gamma(t_2-t_1)}v_y(t_1) + \sqrt{2\gamma}M^{-\frac{1}{2}}\sqrt{\frac{1-e^{2\gamma(t_1-t_2)}}{2\gamma}}\mathcal{N}(0,1)$$
(77)

$$=e^{-\gamma(t_2-t_1)}v_y(t_1) + M^{-\frac{1}{2}}\sqrt{1-e^{2\gamma(t_1-t_2)}}\mathcal{N}(0,1)$$
(78)

The analysis above is performed in one dimension, but readily applies to the  $N \times 3$  dimensional case as the Wiener processes are all independent of each other:

$$v_y(t_2) = e^{-\gamma(t_2 - t_1)} v_y(t_1) + M^{-\frac{1}{2}} \sqrt{1 - e^{2\gamma(t_1 - t_2)}} \mathcal{N}(0, \mathbb{I}_{N \times 3})$$
(79)

1443 Setting  $\Delta t = t_2 - t_1$ , we get the form of the  $\mathcal{O}$  operator (Equation 64) of the BAOAB integrator in 1444 Appendix G.

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# 1447 I HYPERPARAMETERS

All the models trained in this work are SE(3)-equivariant networks

We use a uniform distribution of noise with a range of .01 to .1, There are four hidden layers. Graphs are constructed using bonds as well as neighbors within radius of five. We use ADAM with learning rate .002. For the capped diamines we use 320,000 snapshots for 200 training diamines, whereas for the uncapped we use the timewarp-Large dataset. Models are trained with six parallel processing GPUs and a batch size of forty-two. The model has 8.2 million parameters. For the hidden layers of the network we use 120 scalar and 32 pseudovector features per node and spherical harmonics up to l = 1 The output is 1 vector per node.

For sampling capped diamines, we use  $\sigma = .04$ , whereas for uncapped we use  $\sigma = .06$ . We always use friction of  $\gamma = 0.1$  and a Langevin dynamics step size of  $\delta = \sigma$ .