COHORT SQUEEZE: BEYOND A SINGLE COMMUNICA TION ROUND PER COHORT IN CROSS-DEVICE FEDER ATED LEARNING

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ABSTRACT

Virtually all federated learning (FL) methods, including FedAvg, operate in the following manner: i) an orchestrating server sends the current model parameters to a cohort of clients selected via certain rule, ii) these clients then independently perform a local training procedure (e.g., via SGD or Adam) using their own training data, and iii) the resulting models are shipped to the server for aggregation. This process is repeated until a model of suitable quality is found. A notable feature of these methods is that each cohort is involved in a single communication round with the server only. In this work we challenge this algorithmic design primitive and investigate whether it is possible to "squeeze more juice" out of each cohort than what is possible in a single communication round. Surprisingly, we find that this is indeed the case, and our approach leads to up to 74%reduction in the total communication cost needed to train a FL model in the crossdevice setting. Our method is based on a novel variant of the stochastic proximal point method (SPPM-AS) which supports a large collection of client sampling procedures some of which lead to further gains when compared to classical client selection approaches.

1 INTRODUCTION

031 Federated Learning (FL) is increasingly recognized for its ability to enable collaborative training 032 of a global model across heterogeneous clients, while preserving privacy (McMahan et al., 2016; 033 2017; Kairouz et al., 2019; Li et al., 2020a; Karimireddy et al., 2020b; Mishchenko et al., 2022b; 034 Malinovsky et al., 2024; Yi et al., 2024). This approach is particularly noteworthy in cross-device 035 FL, involving the coordination of millions of mobile devices by a central server for training purposes (Kairouz et al., 2019). This setting is characterized by intermittent connectivity and limited resources. Consequently, only a subset of client devices participates in each communication round. 037 Typically, the server samples a batch of clients (referred to as a *cohort* in FL), and each selected client trains the model received from the server using its local data. Then, the server aggregates the results sent from the selected cohort. Another notable limitation of this approach is the constraint 040 that prevents workers from storing states (operating in a stateless regime), thereby eliminating the 041 possibility of employing variance reduction techniques. We will consider a reformulation of the 042 cross-device objective that assumes a finite number of workers being selected with uniform prob-043 abilities. Given that, in practice, only a finite number of devices is considered, i.e. the following 044 finite-sum objective is considered:

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$$\min_{x \in \mathbb{R}^d} f(x) \coloneqq \frac{1}{n} \sum_{i=1}^n f_i(x).$$
(1)

This reformulation aligns more closely with empirical observations and enhances understanding for illustrative purposes. The extension to the expectation form of the following theory can be found in Appendix F.4.

Current representative approaches in the cross-device setting include FedAvg and FedProx. In our work, we introduce a method by generalizing stochastic proximal point method with arbitray sampling and term as SPPM-AS. This new method is inspired by the stochastic proximal point method

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Figure 1: The total communication cost (defined as TK) with the number of local communication rounds K needed to reach the target accuracy ϵ for the chosen cohort in each global iteration. The dashed red line depicts the communication cost of the FedAvg algorithm. Markers indicate the TKvalue for different learning rates γ of our algorithm SPPM-AS.

(SPPM), a technique notable for its ability to converge under arbitrarily large learning rates and its flexibility in incorporating various solvers to perform proximal steps. This adaptability makes
SPPM highly suitable for cross-device FL (Li et al., 2020a; Yuan & Li, 2022; 2023; Khaled & Jin, 2023; Lin et al., 2024). Additionally, we introduce support for an arbitrary cohort sampling strategy, accompanied by a theoretical analysis. We present novel strategies that include support for client clustering, which demonstrate both theoretical and practical improvements.

Another interesting parameter that allows for control is the number of local communications. Two distinct types of communication, *global* and *local*, are considered. A *global* iteration is defined as a single round of communication between the server and all participating clients. On the other hand, *local* communication rounds are synchronizations that take place within a chosen cohort. Additionally, we introduce the concept of total communication cost, which includes both local and global communication iterations, to measure the overall efficiency of the communication process. The total communication cost naturally depends on several factors. These include the local algorithm used to calculate the prox, the global stepsize, and the sampling technique.

Previous results on cross-device settings consider only one local communication round for the se-084 lected cohort (Li et al., 2020b; Reddi et al., 2020; Li et al., 2020a; Wang et al., 2021a;b; Xu et al., 085 2021; Malinovsky et al., 2023; Jhunjhunwala et al., 2023; Sun et al., 2023; 2024). Our experimental findings reveal that increasing the number of local communication rounds within a chosen cohort 087 per global iteration can indeed lower the total communication cost needed to reach a desired global 088 accuracy level, which we denote as ε . Figure 1 illustrates the relationship between total communication costs and the number of local communication rounds. Assume that the cost of communication 090 per round is 1 unit. K represents the number of local communication rounds per global iteration for 091 the selected cohort, while T signifies the *minimum* number of global iterations needed to achieve the accuracy threshold ϵ . Then, the total cost incurred by our method can be expressed as TK. For 092 comparison, the dashed line in the figure shows the total cost for the FedAvg algorithm, which always sets K to 1, directly equating the number of global iterations to total costs. Our results across 094 various datasets identify the optimal K for each learning rate to achieve ϵ -accuracy. Figure 1 shows that adding more local communication rounds within each global iteration can lead to a significant 096 reduction in the overall communication cost. For example, when the learning rate is set to 1000, the optimal cost is reached with 10 local communication rounds, making K = 10 a more efficient 098 choice compared to a smaller number. On the other hand, at a lower learning rate of 100, the optimal 099 cost of 12 is reached with K = 3. This pattern indicates that as we increase the number of local com-100 munication rounds, the total cost can be reduced, and the optimal number of local communication 101 rounds tends to increase with higher learning rates. 102

Our key *contributions* are summarized as follows:

 We present and analyze SPPM-AS, a novel approach within the stochastic proximal point method framework tailored for cross-device federated learning, which supports arbitrary sampling strategies. Additionally, we provide an analysis of standard sampling techniques and introduce new techniques based on clustering approaches. These novel techniques are theoretically analyzed, offering a thorough comparison between different methods.

 Our numerical experiments, conducted on both convex logistic regression models and non-convex neural networks, demonstrate that the introduced framework enables fine-tuning of parameters to surpass existing state-of-the-art cross-device algorithms. Most notably, we found that increasing the number of local communication rounds within the selected cohort is an effective strategy for reducing the overall communication costs necessary to achieve a specified target accuracy threshold.

• We offer practical guidance on the proper selection of parameters for federated learning applica-tions. Specifically, we examine the potential choices of solvers for proximal operations, considering both convex and non-convex optimization regimes. Our experiments compare first-order and second-order solvers to identify the most effective ones.

METHOD

In this section, we explore efficient stochastic proximal point methods with arbitrary sampling for cross-device FL to optimize the objective (1). Throughout the paper, we denote $[n] \coloneqq \{1, \ldots, n\}$. Our approach builds on the following assumptions.

Assumption 2.1. Function $f_i : \mathbb{R}^d \to \mathbb{R}$ is differentiable for all samples $i \in [n]$.

This implies that the function f is differentiable. The order of differentiation and summation can be interchanged due to the additive property of the gradient operator.

$$\nabla f(x) \stackrel{Eqn. (1)}{=} \nabla \left[\frac{1}{n} \sum_{i=1}^{n} f_i(x) \right] = \frac{1}{n} \sum_{i=1}^{n} \nabla f_i(x)$$

Assumption 2.2. Function $f_i : \mathbb{R}^d \to \mathbb{R}$ is μ -strongly convex for all samples $i \in [n]$, where $\mu > 0$. That is, $f_i(y) + \langle \nabla f_i(y), x - y \rangle + \frac{\mu}{2} ||x - y||^2 \le f_i(x)$, for all $x, y \in \mathbb{R}^d$.

This implies that f is μ -strongly convex and hence has a unique minimizer, which we denote by x_* . We know that $\nabla f(x_{\star}) = 0$. Notably, we do *not* assume f to be L-smooth.

2.1 SAMPLING DISTRIBUTION

Let S be a probability distribution over the 2^n subsets of [n]. Given a random set $S \sim S$, we define $p_i \coloneqq \operatorname{Prob}(i \in S), \quad i \in [n].$

We restrict our attention to proper and nonvacuous random sets.

Assumption 2.3. S is proper (i.e., $p_i > 0$ for all $i \in [n]$) and nonvacuous (i.e., $\operatorname{Prob}(S = \emptyset) = 0$).

Let C be the selected cohort. Given $\emptyset \neq C \subseteq [n]$ and $i \in [n]$, we define

$$v_i(C) := \begin{cases} \frac{1}{p_i} & i \in C \\ 0 & i \notin C \end{cases} \Rightarrow f_C(x) := \frac{1}{n} \sum_{i=1}^n v_i(C) f_i(x) = \sum_{i \in C} \frac{1}{np_i} f_i(x). \tag{2}$$

Note that $v_i(S)$ is a random variable and f_S is a random function. By construction, $E_{S \sim S}[v_i(S)] =$ 1 for all $i \in [n]$, and hence

$$E_{S \sim S}[f_S(x)] = E_{S \sim S}\left[\frac{1}{n}\sum_{i=1}^n v_i(S)f_i(x)\right] = \frac{1}{n}\sum_{i=1}^n E_{S \sim S}[v_i(S)]f_i(x) = \frac{1}{n}\sum_{i=1}^n f_i(x) = f(x).$$

Therefore, the optimization problem in Equation (1) is equivalent to the stochastic optimization problem

$$\min_{x \in \mathbb{R}^d} \left\{ f(x) := \mathcal{E}_{S \sim \mathcal{S}} \left[f_S(x) \right] \right\}.$$
(3)

Further, if for each $C \subset [n]$ we let $p_C := \operatorname{Prob}(S = C)$, then f can be written in the equivalent form

$$f(x) = \mathbb{E}_{S \sim \mathcal{S}} \left[f_S(x) \right] = \sum_{C \subseteq [n]} p_C f_C(x) = \sum_{C \subseteq [n], p_C > 0} p_C f_C(x).$$
(4)

162 2.2 CORE ALGORITHM

Applying SPPM (Khaled & Jin, 2023) to
Equation (3), we arrive at stochastic proximal point method with arbitrary sampling
(SPPM-AS, Algorithm 1):

$$x_{t+1} = \operatorname{prox}_{\gamma f_{S_t}} (x_t)$$

where $S_t \sim S$.

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Algorithm 1 Stochastic Proximal Point Method with Arbitrary Sampling (SPPM-AS)

1: **Input:** starting point $x^0 \in \mathbb{R}^d$, distribution \mathcal{S} over the subsets of [n], learning rate $\gamma > 0$

2: for $t = 0, 1, 2, \dots$ do 3: Sample $S_t \sim S$ 4: $x_{t+1} = \operatorname{prox}_{\gamma f_{S_t}}(x_t)$ 5: end for

Theorem 1 (Convergence of SPPM-AS). Let Assumption 2.1 (differentiability) and Assumption 2.2 (strong convexity) hold. Let S be a sampling satisfying Assumption 2.3, and define

$$\mu_{\rm AS} := \min_{C \subseteq [n], p_C > 0} \sum_{i \in C} \frac{\mu_i}{n p_i}, \quad \sigma_{\star, \rm AS}^2 := \sum_{C \subseteq [n], p_C > 0} p_C \left\| \nabla f_C \left(x_\star \right) \right\|^2.$$
(5)

Let $x_0 \in \mathbb{R}^d$ be an arbitrary starting point. Then for any $t \ge 0$ and any $\gamma > 0$, the iterates of SPPM-AS (Algorithm 1) satisfy

$$\mathbb{E}\left[\|x_{t} - x_{\star}\|^{2}\right] \leq \left(\frac{1}{1 + \gamma\mu_{AS}}\right)^{2t} \|x_{0} - x_{\star}\|^{2} + \frac{\gamma\sigma_{\star,AS}^{2}}{\gamma\mu_{AS}^{2} + 2\mu_{AS}}$$

Theorem interpretation. In the theorem presented above, there are two main terms: $(1/(1+\gamma\mu_{AS}))^{2t}$ and $\gamma\sigma_{\star,AS}^2/(\gamma\mu_{AS}^2+2\mu_{AS})$, which define the convergence speed and neighborhood, respectively. Additionally, there are three hyperparameters to control the behavior: γ (the global learning rate), AS (the sampling type), and T (the number of global iterations). In the following paragraphs, we will explore special cases to provide a clear intuition of how the SPPM-AS theory works.

Interpolation regime. Consider the interpolation regime, characterized by $\sigma_{\star,AS}^2 = 0$. Since we can use arbitrarily large $\gamma > 0$, we obtain an arbitrarily fast convergence rate. Indeed, $(1/(1+\gamma\mu_{AS}))^{2t}$ can be made arbitrarily small for any fixed $t \ge 1$, even t = 1, by choosing γ large enough. However, this is not surprising, since now f and all functions f_{ξ} share a single minimizer, x_{\star} , and hence it is possible to find it by sampling a small batch of functions even a single function f_{ξ} , and minimizing it, which is what the prox does, as long as γ is large enough.

A single step travels far. Observe that for $\gamma = 1/\mu_{AS}$, we have $\gamma \sigma_{\star,AS}^2/(\gamma \mu_{AS}^2 + 2\mu_{AS}) = \sigma_{\star,AS}^2/3\mu_{AS}^2$. In fact, the convergence neighborhood $\gamma \sigma_{\star,AS}^2/(\gamma \mu_{AS}^2 + 2\mu_{AS})$ is bounded above by three times this quantity irrespective of the choice of the stepsize. Indeed, $\frac{\gamma \sigma_{\star,AS}^2}{\gamma \mu_{AS}^2 + 2\mu_{AS}} \leq \min\left\{\frac{\sigma_{\star,AS}^2}{\mu_{AS}^2}, \frac{\gamma \sigma_{\star,AS}^2}{\mu_{AS}}\right\} \leq$ $\frac{\sigma_{\star,AS}^2}{\mu_{AS}^2}$. That means that no matter how far the starting point x_0 is from the optimal solution x_{\star} , if we choose the stepsize γ to be large enough, then we can get a decent-quality solution after a single iteration of SPPM-AS already! Indeed, if we choose γ large enough so that $(1/1 + \gamma \mu_{AS})^2 ||x_0 - x_{\star}||^2 \leq \delta$, where $\delta > 0$ is chosen arbitrarily, then for t = 1 we get $\mathbb{E} \left[||x_1 - x_{\star}||^2 \right] \leq \delta + \sigma_{\star,AS}^2/\mu_{AS}^2$.

Iteration complexity. We have seen above that an accuracy arbitrarily close to (but not reach-208 ing) $\sigma_{\star,AS}^2/\mu_{AS}^2$ can be achieved via a single step of the method, provided that the stepsize γ 209 is large enough. Assume now that we aim for ϵ accuracy, where $\epsilon \leq \sigma_{\star,AS}^2/\mu_{AS}^2$. We can 210 show that with the stepsize $\gamma = \varepsilon \mu_{AS} / \sigma_{\star,AS}^2$, we get $E\left[\|x_t - x_\star\|^2 \right] \leq \varepsilon$ provided that $t \geq \varepsilon$ 211 212 $\left(\frac{\sigma_{\star,AS}^2}{2\varepsilon\mu_{AS}^2} + \frac{1}{2}\right)\log\left(\frac{2\|x_0 - x_\star\|^2}{\varepsilon}\right)$. We provide the proof in Appendix F.5. To ensure thoroughness, we 213 present in Appendix F.9 the lemma of the inexact formulation for SPPM-AS, which offers greater 214 practicality for empirical experimentation. Further insights are provided in the subsequent experi-215 mental section.

General Framework. By allowing the freedom to choose arbitrary algorithms for solving the proximal operator, one can see that SPPM-AS generalizes renowned methods such as FedProx (Li et al., 2020a) and FedAvg (McMahan et al., 2016). In doing so, we are able to adapt our theoret-ical framework to prove the convergence rate of FedProx using a minimalistic set of assumptions, specifically Assumption 2.1 and Assumption 2.2. A more detailed overview of these generalization properties is provided in Appendix B.4.

2.3 ARBITRARY SAMPLING EXAMPLES

Details on simple Full Sampling (FS) and Nonuniform Sampling (NS) are provided in Appendix B.2. In this section, we focus more intently on the sampling strategies that are of particular interest to us.

Nice Sampling (NICE). Choose $\tau \in [n]$ and let S be a random subset of [n] of size τ chosen uniformly at random. Then $p_i = \tau/n$ for all $i \in [n]$. Moreover, let $\binom{n}{\tau}$ represents the number of combinations of n taken τ at a time, $p_C = \frac{1}{\binom{n}{2}}$ whenever $|C| = \tau$ and $p_C = 0$ otherwise. So,

$$\mu_{\rm AS} = \mu_{\rm NICE}(\tau) := \min_{C \subseteq [n], p_C > 0} \sum_{i \in C} \frac{\mu_i}{np_i} = \min_{C \subseteq [n], |C| = \tau} \frac{1}{\tau} \sum_{i \in C} \mu_i,$$

$$\sigma_{\star,\mathrm{AS}}^{2} = \sigma_{\star,\mathrm{NICE}}^{2}(\tau) := \sum_{C \subseteq [n], p_{C} > 0} p_{C} \left\| \nabla f_{C}\left(x_{\star}\right) \right\|^{2} \stackrel{Eqn. (2)}{=} \sum_{C \subseteq [n], |C| = \tau} \frac{1}{\binom{n}{\tau}} \left\| \frac{1}{\tau} \sum_{i \in C} \nabla f_{i}\left(x_{\star}\right) \right\|^{2}$$

It can be shown that $\mu_{\text{NICE}}(\tau)$ is a *nondecreasing* function of τ (Appendix F.6). So, as the minibatch size τ increases, the strong convexity constant $\mu_{\text{NICE}}(\tau)$ can only improve. Since $\mu_{\text{NICE}}(1) =$ $\min_i \mu_i$ and $\mu_{\text{NICE}}(n) = \frac{1}{n} \sum_{i=1}^n \mu_i$, the value of $\mu_{\text{NICE}}(\tau)$ interpolates these two extreme cases as τ varies between 1 and *n*. Conversely, $\sigma_{\star,\text{NICE}}^2(\tau) = \frac{n/\tau - 1}{n-1} \sigma_{\star,\text{NICE}}^2(1)$ is a nonincreasing function, reaching a value of $\sigma^2_{\star,\text{NICE}}(n) = 0$, as explained in Appendix F.6.

Block Sampling (BS). Let C_1, \ldots, C_b be a partition of [n] into b nonempty blocks. For each $i \in [n]$, let B(i) indicate which block *i* belongs to. In other words, $i \in C_j$ if B(i) = j. Let $S = C_j$ with probability $q_j > 0$, where $\sum_j q_j = 1$. Then $p_i = q_{B(i)}$, and hence Equation (5) takes on the form

$$\mu_{\mathrm{AS}} = \mu_{\mathrm{BS}} := \min_{j \in [b]} \frac{1}{nq_j} \sum_{i \in C_j} \mu_i, \quad \sigma_{\star,\mathrm{AS}}^2 = \sigma_{\star,\mathrm{BS}}^2 := \sum_{j \in [b]} q_j \left\| \sum_{i \in C_j} \frac{1}{np_i} \nabla f_i\left(x_\star\right) \right\|^2.$$

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Considering two extreme cases: If b = 1, then SPPM-BS = SPPM-FS = PPM. So, indeed, we recover the same rate as SPPM-FS. If b = n, then SPPM-BS = SPPM-NS. So, indeed, we recover the same rate as SPPM-NS. We provide the detailed analysis in Appendix B.3.

Stratified Sampling (SS). Let C_1, \ldots, C_b be a partition of [n] into b nonempty blocks, as before. For each $i \in [n]$, let B(i) indicate which block does i belong to. In other words, $i \in C_j$ iff B(i) = j. Now, for each $j \in [b]$ pick $\xi_j \in C_j$ uniformly at random, and define $S = \bigcup_{j \in [b]} \{\xi_j\}$. Clearly, $p_i = \frac{1}{|C_{B(i)}|}$. Let's denote $\mathbf{i}_b \coloneqq (i_1, \cdots, i_b)$, $\mathbf{C}_b \coloneqq C_1 \times \cdots \times C_b$. Then, Equation (5) take on the form

$$\mu_{\rm AS} = \mu_{\rm SS} := \min_{\mathbf{i}_b \in \mathbf{C}_b} \sum_{j=1}^b \frac{\mu_{i_j} |C_j|}{n}, \quad \sigma_{\star,\rm AS}^2 = \sigma_{\star,\rm SS}^2 := \sum_{\mathbf{i}_b \in \mathbf{C}_b} \left(\prod_{j=1}^b \frac{1}{|C_j|} \right) \left\| \sum_{j=1}^b \frac{|C_j|}{n} \nabla f_{i_j} \left(x_\star \right) \right\|^2.$$

Considering two extreme cases: If b = 1, then SPPM-SS = SPPM-US. So, indeed, we recover the same rate as SPPM-US. If b = n, then SPPM-SS = SPPM-FS. So, indeed, we recover the same rate as SPPM-FS. We provide the detailed analysis in Appendix B.3.

270 2.4 Comparing Stratified Sampling with Block Sampling and Nice Sampling

Lemma 1 (Stratified Sampling Variance Bounds). Consider the stratified sampling. For each $j \in [b]$, define

$$\sigma_j^2 := \max_{i \in C_j} \left\| \nabla f_i\left(x_\star\right) - \frac{1}{|C_j|} \sum_{l \in C_j} \nabla f_l\left(x_\star\right) \right\|^2$$

In words, σ_j^2 is the maximal squared distance of a gradient (at the optimum) from the mean of the gradients (at optimum) within cluster C_j . Then

$$\sigma_{\star,\mathrm{SS}}^2 \le \frac{b}{n^2} \sum_{j=1}^{b} |C_j|^2 \sigma_j^2 \le b \max\left\{\sigma_1^2, \dots, \sigma_b^2\right\}.$$

Note that Lemma 1 provides insights into how the variance might be reduced through stratified sampling. For instance, in a scenario of complete inter-cluster homogeneity, where $\sigma_j^2 = 0$ for all j, both bounds imply that $0 = \sigma_{\star,SS}^2 \le \sigma_{\star,BS}^2$. Thus, in this scenario, the convergence neighborhood of stratified sampling is better than that of block sampling.

Stratified Sampling Outperforms Block Sampling and Nice Sampling in Convergence Neighborhood. We theoretically compare stratified sampling with block sampling and nice sampling, advocating for stratified sampling as the superior method for future clustering experiments due to its optimal variance properties. We begin with the assumption of *b* clusters of uniform size *b* (Assumption F.10), which simplifies the analysis by enabling comparisons of various sampling methods, all with the same sampling size, *b*: *b*-nice sampling, stratified sampling with *b* clusters, and block sampling where all clusters are of uniform size *b*. Furthermore, we introduce the concept of optimal clustering for stratified sampling (noted as $C_{b,SS}$, Definition F.11) in response to a counterexample where block sampling and nice sampling achieve lower variance than stratified sampling (Appendix F.8). Finally, with Assumption F.10 and Definition F.11 in place, we can compare the convergence neighborhoods of stratified sampling and nice sampling.

Lemma 2. Given Assumption F.10, the following holds: $\sigma_{\star,SS}^2(\mathcal{C}_{b,SS}) \leq \sigma_{\star,NICE}^2$ for arbitrary *b*. Moreover, the variance within the convergence neighborhood of stratified sampling is less than or equal to that of nice sampling: $\frac{\gamma \sigma_{\star,SS}^2}{\gamma \mu_{SS}^2 + 2\mu_{SS}}(\mathcal{C}_{b,SS}) \leq \frac{\gamma \sigma_{\star,NICE}^2}{\gamma \mu_{NICE}^2 + 2\mu_{NICE}}$.

Lemma 2 demonstrates that, under specific conditions, the stratified sampling neighborhood is preferable to that of nice sampling. One might assume that, under the same assumptions, a similar assertion could be made for showing that block sampling is inferior to stratified sampling. However, this has only been theoretically verified for the simplified case where both the block size and the number of blocks are b = 2, as detailed in Appendix F.8.

3 EXPERIMENTS

313 Practical Decision-Making with SPPM-AS. In our analysis of SPPM-AS, guided by theoreti-314 cal foundations of Theorem 1 and empirical evi-315 dence summarized in Table 1, we explore practi-316 cal decision-making for varying scenarios. This 317 includes adjustments in hyperparameters within 318 the framework $KT(\epsilon, \mathcal{S}, \gamma, \mathcal{A}(K))$. Here, ϵ rep-319 resents accuracy goal, S represents the sampling 320 distribution, γ is representing global learning rate 321 (proximal operator parameter), A denotes the 322 proximal optimization algorithm, while K de-323 notes the number of local communication rounds. In table 1,

Table 1:
$$KT(\epsilon, \mathcal{S}, \gamma, \mathcal{A}(K))$$

HP	Control	$KT(\cdots)$	Exp.
γ	$\gamma \uparrow$	$KT\downarrow, \epsilon\uparrow^{(1)}$	D.2
	optimal $(\gamma, K) \uparrow$	\downarrow	3.3
А	μ -convex + BFGS/CG	↓ compared to LocalGD	3.3
	NonCVX and Hier- archical FL + Adam with tuned lr	↓ compared to LocalGD	3.7

⁽¹⁾ ϵ is a convergence neighborhood or accuracy.

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311 312 we summarize how changes on following hyperparameters will influence target metric. With increasing learning rate γ one achieves faster convergence with smaller accuracy, also noted as accuracy-rate tradeoff. Our primary observation that with an increase in both the learning rate, γ , and the number of local steps, K, leads to an improvement in the convergence rate. Employing various local solvers for proximal operators also shows an improvement in the convergence rate compared to FedAvg in both convex and non-convex cases.

331 **Inexact Prox Implementation.** In practice, the proximal operator cannot be calculated exactly, 332 as proposed in the theoretical version of SPPM-AS (Algorithm 1). In our work, we tackle two approaches for estimating the proximal operator. For logistic regression, we use a simplified ap-333 proach that employs a virtual hub for computation. When integrated into a hierarchical FL archi-334 tecture with physical hubs, this approach minimizes communication costs. Standard optimization 335 algorithms such as BFGS and CG are applied to handle the proximal operations. In the neural net-336 work experiments, we use local optimization algorithms to estimate the proximal operator. We treat 337 the argument of the proximal operator as an optimization objective and decompose it into func-338 tions corresponding to each worker: $x_{t+1} = \text{prox}_{\gamma f_{s_t}}(x_t) := \arg \min_y \left(f_{s_t}(y) + \frac{1}{2\gamma} |y - x|^2 \right) = 0$ 339 $\arg\min_{y} \left(\sum_{i \in S_{t}} f_{i}(y) + \frac{1}{2\gamma|S_{t}|} |y - x|^{2} \right) = \arg\min_{y} \left(\sum_{i \in S_{t}} \tilde{f}_{i}(y) \right)$ Thus, various local meth-340 341 ods for minimizing min_y $\left(\sum_{i \in S_t} \tilde{f}_i(y)\right)$ can be applied, as detailed in Appendix A.3. 342

3.1 OBJECTIVE AND DATASETS

Our analysis begins with logistic regression with a convex l_2 regularizer, which can be represented as:

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 $f_i(x) \coloneqq \frac{1}{n_i} \sum_{j=1}^{n_i} \log \left(1 + \exp(-b_{i,j} x^T a_{i,j}) \right) + \frac{\mu}{2} \|x\|^2,$

where μ is the regularization parameter, n_i denotes the total number of data points at client *i*, $a_{i,j}$ are the feature vectors, and $b_{i,j} \in \{-1, 1\}$ are the corresponding labels. Each function f_i exhibits μ -strong convexity and L_i -smoothness, with L_i computed as $\frac{1}{4n_i} \sum_{j=1}^{n_i} ||a_{i,j}||^2 + \mu$. For our experiments, we set μ to 0.1.

Our study utilized datasets from the LibSVM repository (Chang & Lin, 2011), including mushrooms, a6a, ijcnn1.bz2, and a9a. We divided these into feature-wise heterogeneous non-iid splits for FL, detailed in Appendix C.1, with a default cohort size of 10. We primarily examined logistic regression, finding results consistent with our theoretical framework, as discussed extensively in Section 3.3 through Appendix D.2. Additional neural network experiments are detailed in Section 3.7 and Appendix E.

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3.2 ON CHOOSING SAMPLING STRATEGY

365 As shown in Section 2.3, multiple sampling techniques exist. We propose using clustering approach 366 in conjuction with SPPM-SS as the default sampling strategy for all our experiments. The stratified 367 sampling optimal clustering is impractical due to the difficulty in finding x_* ; therefore, we employ 368 a clustering heuristic that aligns with the concept of creating homogeneous worker groups. One 369 such method is K-means, which we use by default. More details on our clustering approach can be 370 found in the Appendix C.1. We compare various sampling techniques in the left panel of Figure 3. 371 Extensive ablations verified the efficiency of stratified sampling over other strategies, due to variance 372 reduction (Lemma 1).

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374 3.3 COMMUNICATION COST REDUCTION THROUGH INCREASED LOCAL COMMUNICATION ROUNDS

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- In this study, we investigate whether increasing the number of local communication rounds, denoted
- 377 In this study, we investigate whether increasing the number of local communication rounds, denoted as *K*, in our proposed algorithm SPPM-SS, can lead to a decrease in the total communication cost



Figure 2: Analysis of total communication costs against local communication rounds for computing the proximal operator. For LocalGD, we align the x-axis to the total local iterations, highlighting the absence of local communication. The aim is to minimize total communication for achieving a predefined global accuracy ϵ , where $||x_T - x_{\star}||^2 < \epsilon$. The optimal step size and minibatch sampling setup for LocalGD are denoted as LocalGD, optim. This showcases a comparison across varying ϵ values and proximal operator solvers (CG and BFGS).



Figure 3: The first column compares sampling methods, while the right two columns analyze convergence relative to popular baselines. $\gamma = 1.0$.

required to converge to a predetermined global accuracy $\epsilon > 0$. In Figure 1, we analyzed vari-ous datasets, including a6a and mushrooms, confirming that higher local communication rounds reduce communication costs, especially with larger learning rates. Our study includes both self-ablation of SPPM-SS across different learning rate scales and comparisons with the widely-used cross-device FL method LocalGD (or FedAvg) on the selected cohort. Ablation studies were conducted with a large empirical learning rate of 0.1, a smaller rate of 0.01, and an optimal rate as discussed by Khaled & Richtárik (2023), alongside minibatch sampling described by Gower et al. (2019).

In Figure 2, we present more extensive ablations. Specifically, we set the base method (Figure 2a) using the dataset a6a, a proximal solver BFGS, and $\epsilon = 5 \cdot 10^{-3}$. In Figure 2b, we explore the use of an alternative solver, CG (Conjugate Gradient), noting some differences in outcomes. For instance, with a learning rate $\gamma = 1000$, the optimal K with CG becomes 7, lower than 10 in the base setting using BFGS. In Figure 2c, we investigate the impact of varying $\epsilon = 10^{-2}$. Our findings consistently show SPPM-SS's significant performance superiority over LocalGD.



Figure 4: The left column shows the Server-hub-client hierarchical FL architecture. For the right two columns: on the left, communication cost for achieving 70% accuracy in hierarchical FL ($c_1 = 0.05$, $c_2 = 1$); on the right, convergence with optimal hyperparameters ($c_1 = 0.05$, $c_2 = 1$).

3.4 Evaluating the Performance of Various Solvers A

We further explore the impact of various solvers on optimizing the proximal operators, showcasing 449 representative methods in Table 2 in the Appendix A.3. A detailed overview and comparison of local 450 optimizers listed in the table are provided in Section A.3, given the extensive range of candidate op-451 tions available. To emphasize key factors, we compare the performance of first-order methods, such 452 as the Conjugate Gradient (CG) method (Hestenes et al., 1952), against second-order methods, like 453 the Broyden-Fletcher-Goldfarb-Shanno (BFGS) algorithm (Broyden, 1967; Shanno, 1970), in the 454 context of strongly convex settings. For non-convex settings, where first-order methods are preva-455 lent in deep learning experiments, we examine an ablation among popular first-order local solvers, 456 specifically choosing MimeLite (Karimireddy et al., 2020a) and FedOpt (Reddi et al., 2020). The comparisons of different solvers for strongly convex settings are presented in Figure 2b, with the 457 non-convex comparison included in the appendix. Upon comparing first-order and second-order 458 solvers in strongly convex settings, we observed that CG outperforms BFGS for our specific prob-459 lem. In neural network experiments, MimeLite-Adam was found to be more effective than FedOpt 460 variations. However, it is important to note that all these solvers are viable options that have led to 461 impressive performance outcomes. 462

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3.5 Comparative Analysis with Baseline Algorithms

465 In this section, we conduct an extensive comparison with several established cross-device FL base-466 line algorithms. Specifically, we examine MB-GD (MiniBatch Gradient Descent with partial client 467 participation), and MB-LocalGD, which is the local gradient descent variant of MB-GD. We default 468 the number of local iterations to 5 and adopt the optimal learning rate as suggested by Gower et al. 469 (2019). To ensure a fair comparison, the cohort size |C| is fixed at 10 for all minibatch methods, including our proposed SPPM-SS. The results of this comparative analysis are depicted in Figure 3. 470 Our findings reveal that SPPM-SS consistently achieves convergence within a significantly smaller 471 neighborhood when compared to the existing baselines. Notably, in contrast to MB-GD and MB-472 LocalGD, SPPM-SS is capable of utilizing arbitrarily large learning rates. This attribute allows for 473 faster convergence, although it does result in a larger neighborhood size. 474

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3.6 HIERARCHICAL FEDERATED LEARNING

477 We extend our analysis to a hub-based hierarchical FL structure, as conceptualized in the left part of 478 Figure 4. This structure envisions a cluster directly connected to m hubs, with each hub m_i serving 479 n_i clients. The clients, grouped based on criteria such as region, communicate exclusively with their 480 respective regional hub, which in turn communicates with the central server. Given the inherent na-481 ture of this hierarchical model, the communication $\cos c_1$ from each client to its hub is consistently 482 lower than the cost c_2 from each hub to the server. We define communication from clients to hubs as local communication and from hubs to the server as global communication. Under SPPM-SS, 483 the total cost is expressed as $(c_1K + c_2)T_{\text{SPPM-SS}}$, while for LocalGD, it is $(c_1 + c_2)T_{\text{LocalGD}}$. As 484 established in Section 3.3, T_{SPPM-SS} demonstrates significant improvement in total communication 485 costs compared to LocalGD within a hierarchical setting. Our objective is to illustrate this by con-

486 trasting the standard FL setting, depicted in Figure 2a with parameters $c_1 = 1$ and $c_2 = 0$, against 487 the hierarchical FL structure, which assumes $c_1 = 0.1$ and $c_2 = 1$, as shown in Figure 2d. Given the 488 variation in c_1 and c_2 values between these settings, a direct comparison of absolute communication 489 costs is impractical. Therefore, our analysis focuses on the ratio of communication cost reduction 490 in comparison to LocalGD. For the base setting, LocalGD's optimal total communication cost is 39 with 12 local iterations, whereas for SPPM-SS ($\gamma = 1000$), it is reduced to 10 with 10 local 491 and 1 global communication rounds, amounting to a 74.36% reduction. With the hierarchical FL 492 structure in Figure 2d, SPPM-SS achieves an even more remarkable communication cost reduction 493 of 94.87%. Further ablation studies on varying local communication cost c_1 in the Appendix D.3 494 corroborate these findings. 495

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3.7 NEURAL NETWORK EVALUATIONS

Our empirical analysis includes experiments on Convolutional Neural Networks (CNNs) using the FEMNIST dataset, as described by Caldas et al. (2018). We designed the experiments to include a total of 100 clients, with each client representing data from a unique user, thereby introducing natural heterogeneity into our study. We employed the Nice sampling strategy with a cohort size of 10. In contrast to logistic regression models, here we utilize training accuracy as a surrogate for the target accuracy ϵ . For the optimization of the proximal operator, we selected the Adam optimizer, with the learning rate meticulously fine-tuned over a linear grid. Detailed descriptions of the training procedures and the CNN architecture are provided in the Appendix E.

506 In the deep learning context, we performed a set of experiments similar to those conducted for the 507 convex case. In Appendix E.2, we review nice, block, and stratified sampling strategies, demon-508 strating the superiority of stratified sampling. Additionally, Appendix E.4 compares various local 509 solvers for the proximal operator. For comparison with the baselines our analysis primarily focuses on the hierarchical FL structure. Initially, we draw a comparison between our proposed method, 510 511 SPPM-AS, and LocalGD. The crux of our investigation is the total communication cost required to achieve a predetermined level of accuracy, with findings detailed in the right part of Figure 4. 512 Significantly, SPPM-AS demonstrates enhanced performance with the integration of multiple local 513 communication rounds. Notably, the optimal number of these rounds tends to increase alongside the 514 parameter γ . For each configuration, the convergence patterns corresponding to the sets of optimally 515 tuned hyperparameters are depicted in Figure 4. 516

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4 CONCLUSION

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Our research challenges the conventional single-round communication model in federated learning
 by presenting a novel approach where cohorts participate in multiple communication rounds. This
 adjustment leads to a significant 74% reduction in communication costs, underscoring the efficacy
 of extending cohort engagement beyond traditional limits. Our method, SPPM-AS, equipped with
 diverse client sampling procedures, contributes substantially to this efficiency. This foundational
 work showcases a pivotal shift in federated learning strategies. Future work could focus on improv ing algorithmic robustness and ensuring privacy compliance.

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⁸¹⁰ A RELATED WORK

812 A.1 CROSS-DEVICE FEDERATED LEARNING

814 This paper delves into the realm of Federated Learning (FL), focusing on the cross-device variant, which presents unique and significant challenges. In FL, two predominant settings are recognized: 815 cross-silo and cross-device scenarios, as detailed in Table 1 of Kairouz et al., 2019. The primary 816 distinction lies in the nature of the clients: cross-silo FL typically involves various organizations 817 holding substantial data, whereas cross-device FL engages a vast array of mobile or IoT devices. In 818 cross-device FL, the complexity is heightened by the inability to maintain a persistent hidden state 819 for each client, unlike in cross-silo environments. This factor renders certain approaches impractical, 820 particularly those reliant on stateful clients participating consistently across all rounds. Given the 821 sheer volume of clients in cross-device FL, formulating and analyzing outcomes in an expectation 822 form is more appropriate, but more complex than in finite-sum scenarios.

823 The pioneering and perhaps most renowned algorithm in cross-device FL is FedAvg (McMahan 824 et al., 2017) and implemented in applications like Google's mobile keyboard (Hard et al., 2018; 825 Yang et al., 2018; Ramaswamy et al., 2019). However, it is noteworthy that popular accelerated 826 training algorithms such as Scaffold (Karimireddy et al., 2020b) and ProxSkip (Mishchenko et al., 827 2022b) are not aligned with our focus due to their reliance on memorizing the hidden state for each client, which is applicable for cross-device FL. Our research pivots on a novel variant within the 829 cross-device framework. Once the cohort are selected for each global communication round, these 830 cohorts engage in what we term as 'local communications' multiple times. The crux of our study is 831 to investigate whether increasing the number of local communication rounds can effectively reduce the total communication cost to converge to a targeted accuracy. 832

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A.2 STOCHASTIC PROXIMAL POINT METHOD

835 Our exploration in this paper centers on the Stochastic Proximal Point Method (SPPM), a method 836 extensively studied for its convergence properties. Initially termed as the incremental proximal point 837 method by Bertsekas (2011), it was shown to converge nonasymptotically under the assumption of 838 Lipschitz continuity for each f_i . Following this, Ryu & Boyd (2016) examined the convergence 839 rates of SPPM, noting its resilience to inaccuracies in learning rate settings, contrasting with the 840 behavior of Stochastic Gradient Descent (SGD). Further developments in SPPM's application were 841 seen in the works of Patrascu & Necoara (2018), who analyzed its effectiveness in constrained op-842 timization, incorporating random projections. Asi & Duchi (2019) expanded the scope of SPPM 843 by studying a generalized method, AProx, providing insights into its stability and convergence rates under convex conditions. The research by Asi et al. (2020) and Chadha et al. (2022) further ex-844 tended these findings, focusing on minibatching and convergence under interpolation in the AProx 845 framework. 846

847 In the realm of federated learning, particularly concerning non-convex optimization, SPPM is also 848 known as FedProx, as discussed in works like those of Li et al. (2020a) and Yuan & Li (2022). 849 However, it is noted that in non-convex scenarios, the performance of FedProx/SPPM in terms of convergence rates does not surpass that of SGD. Beyond federated learning, the versatility of SPPM 850 is evident in its application to matrix and tensor completion such as in the work of Bumin & Huang 851 (2021). Moreover, SPPM has been adapted for efficient implementation in a variety of optimization 852 problems, as shown by Shtoff (2022). While non-convex SPPM analysis presents significant chal-853 lenges, with a full understanding of its convex counterpart still unfolding, recent studies such as the 854 one by Khaled & Jin (2023) have reported enhanced convergence by leveraging second-order sim-855 ilarity. Diverging from this approach, our contribution is the development of an efficient minibatch 856 SPPM method SPPM-AS that shows improved results without depending on such assumptions. Sig-857 nificantly, we also provide the first empirical evidence that increasing local communication rounds 858 in finding the proximal point can lead to a reduction in total communication costs. 859

860 A.3 LOCAL SOLVERS

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In the exploration of local solvers for the SPPM-AS algorithm, the focus is on evaluating the performance impact of various inexact proximal solvers within federated learning settings, spanning both strongly convex and non-convex objectives. Here's a simple summary of the algorithms discussed:

864		Setting	1st order	2nd order	
865		8			
866			Conjugate Gradients (CG)	BFGS	
867		Strongly-Convex		AICN	
868			Sooffnow	LocalNewton	
869			Scalliew		
870			Mime-Adam	Apollo	
871		Nonconvex	FedAdam-AdaGrad	OASIS	
872			FedSpeed	0/1010	
873					
874		Table 2: Local opt	imizers for solving the proxima	al subproblem.	
875					
876					
877	FedAdagrad-Ad	daGrad (Wang et al., 2	2021b): Adapts AdaGrad for bo	oth client and serv	ver sides within
878	federated learn	ing, introducing local	and global corrections to addr	ess optimizer sta	te handling and
879	solution bias.				
880	BFGS (Broyder	n, 1967; Fletcher, 197	0; Goldfarb, 1970; Shanno, 197	0): A quasi-New	ton method that
881	approximates th	he inverse Hessian ma	atrix to improve optimization ef	ficiency, particula	arly effective in
882	strongly conver	x settings but with lin	nitations in distributed impleme	entations.	
883	AICN (Hanzely	vet al 2022). Offers	a global $O(1/k^2)$ convergence	e rate under a se	mi_strong self_
884	concordance as	sumption streamlini	ng Newton's method without the	the need for line se	arches
885	concordance da	sumption, streammin			
886	LocalNewton (Bischoff et al., 2023)	: Enhances local optimization	steps with seco	nd-order infor-
887	mation and glo	bal line search, show	ing efficacy in heterogeneous	data scenarios de	espite a lack of
888	extensive theor	etical grounding.			
889	Fed-LAMB (Ka	rimi et al., 2022): Ex	stends the LAMB optimizer to t	federated settings	s, incorporating
890	layer-wise and	dimension-wise adap	tivity to accelerate deep neural	network training	,•
891	FedSpeed (Su	n et al., 2023): Aims	to overcome non-vanishing bia	ases and client-dr	rift in federated
892	learning through	gh prox-correction ar	ad gradient perturbation steps,	demonstrating e	effectiveness in
893	image classific	ation tasks.		U	
894	Mimo Adam (K	arimiraddy at al 20	20. Witigatas client drift in f	adarated learning	hy integrating
895	global optimize	er states and an SVR	S-style correction term enhance	ring the adaptabil	lity of Adam to
896	distributed setti	ings	-style concetion term, enhance	ang the adaptaon	
897		ings.			
898	OASIS (Jahani	et al., 2021): Utilizes	s local curvature information to	or gradient scalin	g, providing an
899	adaptive, hyper	parameter-light appro	bach that excels in handling ill-	conditioned prob	lems.
900	Apollo (Ma, 20	20): A quasi-Newtor	n method that dynamically inco	orporates curvatu	re information,
901	showing impro	ved efficiency and p	erformance over first-order me	ethods in deep le	arning applica-
902	tions.				
903	Each algorithm	contributes uniquely	to the landscape of local solve	ers in federated le	arning, ranging
904	from enhanced	adaptivity and effici	ency to addressing specific ch	allenges such as	bias, drift, and
905	computational	overhead.	5 I I I I I I I I I I I I I I I I I I I	8	,,
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908	B THEORE	ETICAL OVERVIE	W AND RECOMMENDATI	ONS	
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910	B.1 PARAME	ETER CONTROL			
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912	We have explor	ed the effects of chan	ging the hyperparameters of SP	PM-AS on its the	oretical proper-
913	ties, as summar	rized in Table 3. This	summary shows that as the lea	rning rate increas	ses, the number
914	of iterations red	quired to achieve a ta	rget accuracy decreases, thoug	h this comes with	an increase in
915	neighborhood s	size. Focusing on sai	npling strategies, for SPPM-N	i∪⊨ employing N	HCE sampling,
916	an increase in t	The sampling size τ_S is that stratified sources	esuits in lewer iterations (T) at	nu a smaller neig	noornood. Fur-

917 thermore, given that stratified sampling outperforms both block sampling and NICE sampling, we recommend adopting stratified sampling, as advised by Lemma 1.

Table 3: Theoretical summary

Hyperparameter	Control	Rate (T)	Neighborhood
γ	↑	\downarrow	↑
S	$\tau_{\mathcal{S}}\uparrow^{(1)}$	\downarrow	\downarrow
U	Stratified sampling optimal clustering instead of BS or NICE sampling	\downarrow	Lemma 1

⁽¹⁾ We define $\tau_{\mathcal{S}} := \mathbb{E}_{S \sim \mathcal{S}} [|S|]$.

F 1 1 4	A 1 *.	1.	•
Table 4.	Arhitrary	samplings	comparison
	7 monual y	samprings	companson

Setting/Requirement	$\mu_{ m AS}$	$\sigma_{\star,\mathrm{AS}}$
Full	$\frac{1}{n}\sum_{i=1}^{n}\mu_i$	0
Non-Uniform	$\min_i \frac{\mu_i}{np_i}$	$\frac{1}{n}\sum_{i=1}^{n}\frac{1}{np_{i}}\left\ \nabla f_{i}\left(x_{\star}\right)\right\ ^{2}$
Nice	$\min_{C \subseteq [n], C = \tau} \frac{1}{\tau} \sum_{i \in C} \mu_i$	$\sum_{C \subseteq [n], C = \tau} \frac{1}{\binom{n}{\tau}} \left\ \frac{1}{\tau} \sum_{i \in C} \nabla f_i \left(x_\star \right) \right\ ^2$
Block	$\min_{j \in [b]} \frac{1}{nq_j} \sum_{i \in C_j} \mu_i$	$\sum_{j \in [b]} q_j \left\ \sum_{i \in C_j} \frac{1}{np_i} \nabla f_i \left(x_\star \right) \right\ ^2$
Stratified	$\min_{\mathbf{i}_b \in \mathbf{C}_b} \sum_{j=1}^b \frac{\mu_{i_j} C_j }{n}$	$\sum_{\mathbf{i}_{b} \in \mathbf{C}_{b}} \left(\prod_{j=1}^{b} \frac{1}{ C_{j} } \right) \left\ \sum_{j=1}^{b} \frac{ C_{j} }{n} \nabla f_{i_{j}}\left(x_{\star}\right) \right\ ^{2}$
Statilled		Upper bound: $\frac{b}{n^2} \sum_{j=1}^{b} C_j ^2 \sigma_j^2$

B.2 COMPARISON OF SAMPLING STRATEGIES

Full Sampling (FS). Let S = [n] with probability 1. Then SPPM-AS applied to Equation (9) becomes PPM (Moreau, 1965; Martinet, 1970) for minimizing f. Moreover, in this case, we have $p_i = 1$ for all $i \in [n]$ and Equation (5) takes on the form

$$\mu_{\rm AS} = \mu_{\rm FS} := \frac{1}{n} \sum_{i=1}^{n} \mu_i, \quad \sigma_{\star,\rm AS}^2 = \sigma_{\star,\rm FS}^2 := 0.$$

Note that μ_{FS} is the strong convexity constant of f, and that the neighborhood size is zero, as we would expect.

Nonuniform Sampling (NS). Let $S = \{i\}$ with probability $p_i > 0$, where $\sum_i p_i = 1$. Then Equation (5) takes on the form

$$\mu_{\rm AS} = \mu_{\rm NS} := \min_{i} \frac{\mu_{i}}{np_{i}}, \quad \sigma_{\star,\rm AS}^{2} = \sigma_{\star,\rm NS}^{2} := \frac{1}{n} \sum_{i=1}^{n} \frac{1}{np_{i}} \left\| \nabla f_{i} \left(x_{\star} \right) \right\|^{2}$$

958 If we take $p_i = \frac{\mu_i}{\sum_{j=1}^n \mu_j}$ for all $i \in [n]$, we shall refer to Algorithm 1 as SPPM with importance 959 sampling (SPPM-IS). In this case,

$$\mu_{\rm NS} = \mu_{\rm IS} := \frac{1}{n} \sum_{i=1}^{n} \mu_i, \quad \sigma_{\star,\rm NS}^2 = \sigma_{\star,\rm IS}^2 := \frac{\sum_{i=1}^{n} \mu_i}{n} \sum_{i=1}^{n} \frac{\|\nabla f_i(x_\star)\|^2}{n\mu_i}.$$

This choice maximizes the value of μ_{NS} (and hence minimizes the first part of the convergence rate) over the choice of the probabilities.

Table 4 summarizes the parameters associated with various sampling strategies, serving as a concise overview of the methodologies discussed in the main text. This summary facilitates a quick comparison and reference.

970 B.3 EXTREME CASES OF BLOCK SAMPLING AND STRATIFIED SAMPLING

Extreme cases of block sampling. We now consider two extreme cases:

• If b = 1, then SPPM-BS = SPPM-FS = PPM. Let's see, as a sanity check, whether we recover the right rate as well. We have $q_1 = 1, C_1 = [n], p_i = 1$ for all $i \in [n]$, and the expressions for μ_{AS} and $\sigma_{\star, BS}^2$ simplify to

$$\mu_{\rm BS} = \mu_{\rm FS} := \frac{1}{n} \sum_{i=1}^{n} \mu_i, \sigma_{\star,\rm BS}^2 = \sigma_{\star,\rm FS}^2 := 0.$$

So, indeed, we recover the same rate as SPPM-FS.

• If b = n, then SPPM-BS = SPPM-NS. Let's see, as a sanity check, whether we recover the right rate as well. We have $C_i = \{i\}$ and $q_i = p_i$ for all $i \in [n]$, and the expressions for μ_{AS} and $\sigma_{\star,BS}^2$ simplify to

$$\mu_{\rm BS} = \mu_{\rm NS} := \min_{i \in [n]} \frac{\mu_i}{np_i}, \quad \sigma_{\star,\rm BS}^2 = \sigma_{\star,\rm NS}^2 := \frac{1}{n} \sum_{i=1}^n \frac{1}{np_i} \|\nabla f_i(x_\star)\|^2.$$

So, indeed, we recover the same rate as SPPM-NS.

Extreme cases of stratified sampling. We now consider two extreme cases:

• If b = 1, then SPPM-SS = SPPM-US. Let's see, as a sanity check, whether we recover the right rate as well. We have $C_1 = [n], |C_1| = n, \left(\prod_{j=1}^b \frac{1}{|C_j|}\right) = \frac{1}{n}$ and hence

$$\mu_{\rm SS} = \mu_{\rm US} := \min_{i} \mu_{i}, \quad \sigma_{\star,\rm SS}^{2} = \sigma_{\star,\rm US}^{2} := \frac{1}{n} \sum_{i=1}^{n} \|\nabla f_{i}(x_{\star})\|^{2}.$$

So, indeed, we recover the same rate as SPPM-US.

• If b = n, then SPPM-SS = SPPM-FS. Let's see, as a sanity check, whether we recover the right rate as well. We have $C_i = \{i\}$ for all $i \in [n], \left(\prod_{j=1}^{b} \frac{1}{|C_i|}\right) = 1$, and hence

$$\mu_{\rm SS} = \mu_{\rm FS} := \frac{1}{n} \sum_{i=1}^{n} \mu_i, \quad \sigma_{\star,\rm SS}^2 = \sigma_{\star,\rm FS}^2 := 0.$$

So, indeed, we recover the same rate as SPPM-FS.

1015 B.4 FEDERATED AVERAGING SPPM BASELINES

In this section we propose two new algorithms based on federated averaging principle. Since to the best of our knowledge there are no federated averaging analyses within the same assumptions, we provide analysis of modified versions of SPPM-AS.

Averaging on $\operatorname{prox}_{\gamma f_i}$. We introduce FedProx-SPPM-AS (see Algorithm 2), which is inspired by the principles of FedProx (Li et al., 2020a). Unlike the SPPM-AS approach where a proximal operator is computed for the chosen cohort as a whole, in FedProx-SPPM-AS, we compute and then average the proximal operators calculated for each member within the cohort. One can see, that the FedProx is the simple case of this algorithm, when number of local communication rounds K = 1.

1027 1028	Algorithm 2 Provinal Averaging SPPM_AS	Algorithm 3 Federated Averaging SPPM-AS (FedAvg-SPPM-AS)
1029 1030 1031	(FedProx-SPPM-AS) 1: Input: starting point $x_{0,0} \in \mathbb{R}^d$, arbitrary	1: Input: starting point $x_{0,0} \in \mathbb{R}^d$, arbitrary sampling distribution S , global learning rate
1032	sampling distribution S , learning rate $\gamma > 0$, local communication rounds K .	$\gamma > 0$, local learning rate $\alpha > 0$, local com- munication rounds K
1034	2: for $t = 0, 1, 2, \dots, T-1$ do 3: Sample $S_t \sim S$	2: for $t = 0, 1, 2, \cdots, T-1$ do 3: Sample $S_t \sim S$
1035 1036	4: for $k = 0, 1, 2, \dots K - 1$ do 5: $x_{k+1,t} = \sum_{i \in S_t} \frac{1}{ S_i } \operatorname{prox}_{\gamma f_i}(x_{k,t})$	4: $\forall i \in S_t \ f_{i,t}(x) \leftarrow f_i(x) + \frac{1}{2\gamma} \ x - x_t\ ^2$ 5: for $k = 0, 1, 2, \dots K - 1$ do
1037 1038	6: end for 7: $x_{0,t+1} \leftarrow x_{K,t}$	6: $x_{k+1,t} = \sum_{i \in S_t} \frac{1}{ S_t } \operatorname{prox}_{\alpha \tilde{f}_{i,t}}(x_{k,t})$ 7: end for
1039 1040	8: end for 9: Output: $x_0 T$	8: $x_{0,t+1} \leftarrow x_{K,t}$
1041		10: Output: $x_{0,T}$

Here, we employ a proof technique similar to that of Theorem 1 and obtain the following convergence.

Theorem 2 (FedProx convergence). Let the number of local iterations K = 1, and assume that Assumption 2.1 (differentiability) and Assumption 2.2 (strong convexity) hold. Let $x_0 \in \mathbb{R}^d$ be an arbitrary starting point. Then, for any $t \ge 0$ and any $\gamma > 0$, the iterates of FedProx-SPPM (as described in Algorithm 2) satisfy:

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Where
$$A_{\mathcal{S}} \coloneqq \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{1}{1 + \gamma \mu_i} \right]$$
 and $B_{\mathcal{S}} \coloneqq \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{\gamma}{(1 + \gamma \mu_i)\mu_i} \|\nabla f_i(x_\star))\|^2 \right].$

Compared to the theoretical analysis for convex functions provided in Li et al. (2020a), our theoretical bound does not rely on the "gradient boundedness" assumption or the L-smooth constant.

Federated averaging for prox approximation. An alternative method involves estimating the proximal operator by averaging the proximal operators calculated for each worker's function. We call it *Federated Averaging Stochastic Proximal Point Method* (FedAvg-SPPM-AS, see Algorithm 3).
 (FedAvg-SPPM-AS, see Algorithm 3).

After selecting and fixing a sample of workers S_k , the main objective is to calculate the proximal operator. This can be accomplished by approximating the proximal calculation with the goal of minimizing $\tilde{f}_S(x) = f_S(x) + \frac{2}{\gamma} ||x - x_t||^2$. Essentially, this method utilizes FedProx as a local solver for computing the proximal operator. It can be observed that this approach is equivalent to FedProx-SPPM-AS, as at each local step we calculate

$$\begin{array}{l} 1067 \\ 1068 \\ 1069 \end{array} \quad \operatorname{prox}_{\alpha \tilde{f}_i}(x_{k,t}) \coloneqq \operatorname*{arg\,min}_{z \in \mathbb{R}^d} \left[\tilde{f}_i(z) + \frac{2}{\alpha} \|z - x_{k,t}\|^2 \right] = \operatorname*{arg\,min}_{z \in \mathbb{R}^d} \left[f_i(z) + \left(\frac{2}{\gamma} + \frac{2}{\alpha} \right) \|z - x_{k,t}\|^2 \right].$$

1070 It follows that FedProx is equivalent to FedAvg-SPPM-AS when the number of communication 1071 rounds is set to K = 1. Thus, we can conclude that FedProx is a specific instance of SPPM-AS.

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- ¹⁰⁷³ C TRAINING DETAILS
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1075 C.1 NON-IID DATA GENERATION

In our study, we validate performance and compare the benefits of SPPM-AS over SPPM using
 well-known datasets such as mushrooms, a6a, w6a, and ijcnn1.bz2 from LibSVM (Chang
 & Lin, 2011). To ensure relevance to our research focus, we adopt a feature-wise non-IID setting, characterized by variation in feature distribution across clients. This variation is introduced by



Figure 5: t-SNE visualization of cluster-features across data samples on clients.

clustering the features using the K-means algorithm, with the number of clusters set to 10 and the
number of clients per cluster fixed at 10 for simplicity. We visualize the clustered data using t-SNE
in Figure 5, where we observe that the data are divided into 10 distinct clusters with significantly
spaced cluster centers.

1099 C.2 SAMPLING

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To simulate random sampling among clients within these 10 clusters, where each cluster comprises 10 clients, we consider two contrasting scenarios:

- Case I SPPM-BS: Assuming clients within the same cluster share similar features and data distributions, sampling all clients from one cluster (i.e., C = 10 clients) results in a homogeneous sample.
 - *Case II* SPPM-SS: Conversely, by traversing all 10 clusters and randomly sampling one client from each, we obtain a group of 10 clients representing maximum heterogeneity.

1109 We hypothesize that any random sampling from the 100 clients will yield performance metrics lying 1110 between these two scenarios. In Figure 6, we examine the impact of sampling clients with varying degrees of heterogeneity using a fixed learning rate of 0.1. Our findings indicate that heterogeneous 1111 sampling results in a significantly smaller convergence neighborhood σ_{\star}^2 . This outcome is attributed 1112 to the broader global information captured through heterogeneous sampling, in contrast to homo-1113 geneous sampling, which increases the data volume without contributing additional global insights. 1114 As these two sampling strategies represent the extremes of arbitrary sampling, any random selection 1115 will fall between them in terms of performance. Given their equal cost and the superior performance 1116 of the SPPM-SS strategy in heterogeneous FL environments, we designate SPPM-SS as our default 1117 sampling approach. 1118

1119 C.3 SPPM-AS ALGORITHM ADAPTATION FOR FEDERATED LEARNING

In the main text, Algorithm 1 outlines the general form of SPPM-AS. For the convenience of implementation in FL contexts and to facilitate a better understanding, we introduce a tailored version of the SPPM-AS algorithm specific to FL, designated as Algorithm 4. Notably, as stratified sampling is adopted as our default method, this adaptation of the algorithm specifically addresses the nuances of the block sampling approach. We also conducted arbitrary sampling on synthetic datasets and neural networks to demonstrate the algorithm's versatility.

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¹¹⁸⁸ D Additional Experiments on Logistic Regression

1190 D.1 COMMUNICATION COST ON VARIOUS DATASETS TO A TARGET ACCURACY

In Figure 1, we presented the total communication cost relative to the number of rounds required to achieve the target accuracy for the selected cohort. In this section, we provide more details on how is this figure was obtained and present additional results for various datasets.



Figure 7: Total communication cost with respect to the local communication round. For LocalGD, K represents the local communication round K for finding the prox of the current model. For LocalGD, we slightly abuse the x-axis, which represents the total number of local iterations, no local communication is required. We calculate the total communication cost to reach a fixed global accuracy ϵ such that $||x_t - x_*||^2 < \epsilon$. LocalGD, optim represents using the theoretical optimal stepsize of LocalGD with minibatch sampling.



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D.2 CONVERGENCE SPEED AND $\sigma^2_{\star SS}$ TRADE-OFF

1233 Unlike SGD-type methods such as MB-GD and MB-LocalGD, in which the largest allowed learn-1234 ing rate is 1/A, where A is a constant proportion to the smoothness of the function we want to 1235 optimize (Gower et al., 2019). For larger learning rate, SGD-type method may not converge and 1236 exploding. However, for stochastic proximal point methods, they have a very descent benefit of allowing arbitrary learning rate. In this section, we verify whether our proposed method can allow 1237 arbitrary learning rate and whether we can find something interesting. We considered different learn-1238 ing rate scale from 1e-5 to 1e+5. We randomly selected three learning rates [0.1, 1, 100] for visual 1239 representation with the results presented in Figure 8 and Figure 9. We found that a larger learning 1240 rate leads to a faster convergence rate but results in a much larger neighborhood, $\sigma_{\star,SS}^2/\mu_{SS}^2$. This 1241 can be considered a trade-off between convergence speed and neighborhood size, $\sigma_{\star SS}^2$. By default,

we consider setting the learning rate to 1.0 which has a good balance between the convergence speed and the neighborhood size.

In this section, we extend our analysis by providing additional results across a broader range of datasets and varying learning rates. Specifically, Figure 8 illustrates the outcomes using 4 local communication rounds (K = 4), while Figure 9 details the results for 16 local communication rounds (K = 16). Previously, in Figure 1, we explored the advantages of larger K values. Here, our focus shifts to determining if similar trends are observable across different K values. Through comprehensive evaluations on various datasets and multiple K settings, we have confirmed that lower learning rates in SPPM-AS result in slower convergence speeds; however, they also lead to a smaller final convergence neighborhood.

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D.3 ADDITIONAL EXPERIMENTS ON HIERARCHICAL FEDERATED LEARNING

1255 In Figure 2d of the main text, we detail the total communication cost for hierarchical Federated Learning (FL) utilizing parameters $c_1 = 0.1$ and $c_2 = 1$ on the a 6a dataset. Our findings re-1256 veal that SPPM-AS achieves a significant reduction in communication costs, amounting to 94.87%, 1257 compared with the conventional FL setting where $c_1 = 1$ and $c_2 = 1$, which shows a 74.36% 1258 reduction. In this section, we extend our analysis with comprehensive evaluations on additional 1259 datasets, namely ijcnn1.bz2, a9a, and mushrooms. Beyond considering $c_1 = 0.1$, we further 1260 explore the impact of reducing the local communication cost from each client to the corresponding 1261 hub to $c_1 = 0.05$. The results, presented in Figure 10 and the continued Figure 11, reinforce our 1262 observation: hierarchical FL consistently leads to further reductions in communication costs. A 1263 lower c_1 parameter correlates with even greater savings in communication overhead. These results 1264 not only align with our expectations but also underscore the efficacy of our proposed SPPM-AS in 1265 cross-device FL settings.



Figure 10: The total communication cost is analyzed with respect to the number of local communication rounds. For LocalGD, K represents the local communication round used for finding the prox of the current model. In the case of LocalGD, we slightly abuse the x-axis to represent the total number of local iterations, as no local communication is required. We calculate the total communication cost needed to reach a fixed global accuracy ϵ , such that $||x_t - x_*||^2 < \epsilon$. LocalGD, optim denotes the use of the theoretically optimal stepsize for LocalGD with minibatch sampling. Comparisons are made between different prox solvers (CG and BFGS).

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E ADDITIONAL NEURAL NETWORK EXPERIMENTS

1291 E.1 EXPERIMENT DETAILS

For our neural network experiments, we used the FEMNIST dataset (Caldas et al., 2018). Each client was created by uniformly selecting from user from original dataset, inherently introducing heterogeneity among clients. We tracked and reported key evaluation metrics—training and testing loss and accuracy—after every 5 global communication rounds. The test dataset was prepared by



Figure 11: Total communication cost with respect to the local communication round.

Layer	Output Shape	# of Trainable Parameters	Activation	Hyperparameters
Input	(28, 28, 1)	0		
Conv2d	(24, 24, 32)	832	ReLU	kernel size = 5 ; strides = $(1, 1)$
Conv2d	$(10 \ 10 \ 64)$	51 264	ReLU	kernel size = 5
conv2d	(10, 10, 01)	51,201	Relo	strides = $(1, 1)$
MaxPool2d	(5, 5, 64)	0		pool size = $(2, 2)$
Flatten	6400	Ő		poor size (2, 2)
Dense	128	819.328	ReLU	
Dense	62	7.998	softmax	

Table 5: Architecture of the CNN model for FEMNIST symbol recognition.

dividing each user's data into a 9:1 ratio, following the partitioning approach of the FedLab framework (Zeng et al., 2023). For the SPPM-AS algorithm, we selected Adam as the optimizer for the
proximal operator. The learning rate was determined through a grid search across the following
range: [0.0001, 0.0005, 0.001, 0.005, 0.01, 0.05, 0.1, 0.5]. The model architecture comprises a convolutional neural network (CNN) with the following layers: Conv2d(1, 32, 5), ReLU, Conv2d(32,
64, 5), MaxPool2d(2, 2), a fully connected (FC) layer with 128 units, ReLU, and another FC layer
with 128 units, as specified in Table 5. Dropout, learning rate scheduling, gradient clipping, etc.,
were not used to improve the interpretability of results.

We explore various values of targeted training accuracy, as illustrated in Figure 12. This analysis helps us understand the impact of different accuracy thresholds on the model's performance. For instance, we observe that as the target accuracy changes, SPPM-NICE consistently outperforms LocalGD in terms of total communication cost. As the target accuracy increases, the performance gap between these two algorithms also widens. Additionally, we perform ablation studies on different values of c_1 , as shown in Figure 13, to assess their effects on the learning process. Here, we note that with $c_2 = 0.2$, SPPM-NICE performs similarly to LocalGD, suggesting that an increase in c_2 value could narrow the performance gap between SPPM-NICE and LocalGD.



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Figure 12: Varying targeted training accuracy level for SPPM-AS.

E.2 SAMPLINGS DEEP LEARNING PRACTICAL COMPARISON

1396 In the arbitrary samplings theory section 2.3, we referred to several sampling techniques such as batch sampling and nice sampling, which require dividing workers into clusters. We have reviewed 1398 the theoretical properties of clustering using the values of gradients at the optimum x_* . However, in 1399 practice, this information is unavailable, necessitating the use of various clustering methods based 1400 on heuristics. To achieve this, we can apply unsupervised algorithms such as feature-wise K-means. In the federated learning scenario, each worker's dataset consists of numerous data points, so we 1401 need to create some representation of them. One of the most straightforward methods is to average 1402 the dataset points. The 2-dimensional t-SNE representation of the provided clustering is shown in 1403 Figure 14. Cluster patterns are visible, but they are not as well separated as theory suggests.



1457 SGD as the worker optimizer, FedAdam - the Federated Averaging algorithm with Adam as the worker optimizer, FedAdam-Adam based on the FedOpt framework (Reddi et al., 2020), and finally



Figure 15: Comparison of stratified, block, and nice samplings based on training data convergence. For stratified and block sampling, feature-wise averaged K-Means clustering into 10 clusters is used. All parameters, aside from samplings, are fixed at the same values: the number of local communication rounds (T) is 3, the number of worker training epochs is 3, and the number of sampled workers is 10.



Figure 16: Different local solvers for prox baselines for training a CNN model over 100 workers using data from the FEMNIST dataset. The number of local communication rounds is fixed at 3 and the number of worker optimizer steps is fixed at 3. Nice sampling with a minibatch size of 10 is used. γ is fixed at 1.0.

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¹⁵⁶⁶ F MISSING PROOF AND ADDITIONAL THEORETICAL ANALYSIS

1568 F.1 FACTS USED IN THE PROOF 1569

1570 Fact F.1 (Differentiation of integral with a parameter (theorem 2.27 from Folland (1984))). Suppose 1571 that $f: X \times [a,b] \to \mathbb{C}(-\infty < a < b < \infty)$ and that $f(\cdot,t): X \to \mathbb{C}$ is integrable for each 1572 $t \in [a,b]$. Let $F(t) = \int_X f(x,t)d\mu(x)$.

[label=.]

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1614 1615 1. Suppose that there exists $g \in L^1(\mu)$ such that $|f(x,t)| \leq g(x)$ for all x, t. If $\lim_{t \to t_0} f(x,t) = f(x,t_0)$ for every x, then $\lim_{t \to t_0} F(t) = F(t_0)$; in particular, if $f(x, \cdot)$ is continuous for each x, then F is continuous.

2. Suppose that $\partial f/\partial t$ exists and there is a $g \in L^1(\mu)$ such that $|(\partial f/\partial t)(x,t)| \leq g(x)$ for all x, t. Then F is differentiable and $F'(x) = \int (\partial f/\partial t)(x,t)d\mu(x)$.

Fact F.2 (Tower Property). For any random variables X and Y, we have

$$\mathbb{E}\left[\mathbb{E}\left[X|Y\right]\right] = \mathbb{E}\left[X\right]$$

Fact F.3 (Every point is a fixed point (Khaled & Jin, 2023)). Let $\varphi : \mathbb{R}^d \to \mathbb{R}$ be a convex differentiable function. Then

$$\operatorname{prox}_{\gamma\omega}(x + \gamma \nabla \varphi(x)) = x, \qquad \forall \gamma > 0, \quad \forall x \in \mathbb{R}^d.$$

1588 In particular, if x_{\star} is a minimizer of φ , then $\operatorname{prox}_{\gamma\varphi}(x_{\star}) = x_{\star}$.

¹⁵⁹⁰ *Proof.* Evaluating the proximity operator is equivalent to

$$\operatorname{prox}_{\gamma\varphi}(y) = \operatorname*{arg\,min}_{x \in \mathbb{R}^d} \left(\varphi(x) + \frac{1}{2\gamma} \|x - y\|^2 \right).$$

This is a strongly convex minimization problem for any $\gamma > 0$, hence the (necessarily unique) minimizer $x = \operatorname{prox}_{\gamma\varphi}(y)$ of this problem satisfies the first-order optimality condition

$$abla arphi(x) + rac{1}{\gamma}(x-y) = 0.$$

Solving for y, we observe that this holds for $y = x + \gamma \nabla \phi(x)$. Therefore, $x = \operatorname{prox}_{\gamma \varphi}(x + \gamma \nabla \phi(x))$.

1602 Fact F.4 (Contractivity of the prox (Mishchenko et al., 2022a)). If φ is differentiable and μ -strongly convex, then for all $\gamma > 0$ and for any $x, y \in \mathbb{R}^d$ we have

$$\left\|\operatorname{prox}_{\gamma\varphi}(x) - \operatorname{prox}_{\gamma\varphi}(y)\right\|^2 \le \frac{1}{(1+\gamma\mu)^2} \left\|x - y\right\|^2.$$

Fact F.5 (Recurrence (Khaled & Jin, 2023, Lemma 1)). Assume that a sequence $\{s_t\}_{t\geq 0}$ of positive real numbers for all $t \geq 0$ satisfies

$$s_{t+1} \leq as_t + b$$
,

where 0 < a < 1 and $b \ge 0$. Then the sequence for all $t \ge 0$ satisfies

$$s_t \le a^t s_0 + b \min\left\{t, \frac{1}{1-a}\right\}.$$

1616 *Proof.* Unrolling the recurrence, we get

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$$s_t \le as_{t-1} + b \le a(as_{t-2} + b) + b \le \dots \le a^t s_0 + b \sum_{i=0}^{t-1} a^i.$$

We can now bound the sum $\sum_{i=0}^{t-1} a^i$ in two different ways. First, since a < 1, we get the estimate

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$$\sum_{i=0}^{t-1} a^i \le \sum_{i=0}^{t-1} 1 = t.$$

1625 Second, we sum a geometic series

$$\sum_{i=0}^{t-1} a^i \le \sum_{i=0}^{\inf} a^i = \frac{1}{1-a}.$$

1630 Note that either of these bounds can be better. So, we apply the best of these bounds. Substituing the above two bounds gived the target inequality. \Box

1633 F.2 SIMPLIFIED PROOF OF SPPM

 $||x_{t+1} - x_{\star}||^2 = \left\| \operatorname{prox}_{\gamma f_{\xi_{\star}}} - x_{\star} \right\|^2$

1635 We provide a simplified proof of SPPM (Khaled & Jin, 2023) in this section. Using the fact that 1636 $x_{\star} = \operatorname{prox}_{\gamma f_{\xi_t}}(x_{\star} + \gamma \nabla f_{\xi_t}(x_{\star}))$ (see Fact F.3) and then applying contraction of the prox (Fact F.4), 1637 we get

 $\overset{(Fact\ F.3)}{=} \left\| \operatorname{prox}_{\gamma f_{\xi_t}}(x_t) - \operatorname{prox}_{\gamma f_{\xi_t}}(x_\star + \gamma \nabla f_{\xi_t}(x_\star)) \right\|^2$ $\overset{(Fact\ F.4)}{\leq} \frac{1}{(1+\gamma\mu)^2} \|x_t - (x_\star + \gamma \nabla f_{\xi_t}(x_\star))\|^2$ $= \frac{1}{(1+\gamma\mu)^2} \left(\|x_t - x_\star\|^2 - 2\gamma \langle \nabla f_{\xi_t}(x_\star), x_t - x_\star \rangle + \gamma^2 \| \nabla f_{\xi_t}(x_\star) \|^2 \right).$

Taking expectation on both sides, conditioned on x_t , we get

$$\mathbb{E}\left[\|x_{t+1} - x_{\star}\|^{2} | x_{t}\right] \leq \frac{1}{(1+\gamma\mu)^{2}} \left(\|x_{t} - x_{\star}\|^{2} - 2\gamma \left\langle \mathbb{E}\left[\nabla f_{\xi_{t}}(x_{\star})\right], x_{t} - x_{\star}\right\rangle + \gamma^{2} \mathbb{E}\left[\|\nabla f_{\xi_{t}}(x_{\star})\|^{2}\right]\right)$$
$$= \frac{1}{(1+\gamma\mu)^{2}} \left(\|x_{t} - x_{\star}\|^{2} + \gamma^{2}\sigma_{\star}^{2}\right),$$

where we used the fact that $\mathbb{E}\left[\nabla f_{\xi_t}(x_\star)\right] = \nabla f(x_\star) = 0$ and $\sigma_\star^2 := \mathbb{E}\left[\left\|\nabla f_{\xi_t}(x_\star)\right\|^2\right]$. Taking expectation again and applying the tower property (Fact F.2), we get

$$\mathbb{E}\left[\|x_{t+1} - x_{\star}\|^{2}\right] \leq \frac{1}{(1+\gamma\mu)^{2}} \left(\|x_{t} - x_{\star}\|^{2} + \gamma^{2}\sigma_{\star}^{2}\right).$$

1664 It only remains to solve the above recursion. Luckily, that is exactly what Fact F.5 does. In particular, 1665 we use it with $s_t = \mathbb{E}\left[\|x_t - x_\star\|^2\right], a = \frac{1}{(1+\gamma\mu)^2}$ and $b = \frac{\gamma^2 \sigma_\star^2}{(1+\gamma\mu)^2}$ to get

$$\mathbb{E}\left[\|x_{t} - x_{\star}\|^{2}\right] \stackrel{(Fact \ F.5)}{\leq} \left(\frac{1}{1+\gamma\mu}\right)^{2t} \|x_{0} - x_{\star}\|^{2} + \frac{\gamma^{2}\sigma_{\star}^{2}}{(1+\gamma\mu)^{2}} \min\left\{t, \frac{(1+\gamma\mu)^{2}}{(1+\gamma\mu)^{2}-1}\right\}$$

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$$\leq \left(\frac{1}{1+\gamma\mu}\right) \|x_0 - x_\star\|^2 + \frac{1}{(1+\gamma\mu)^2 - 1}$$

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$$\leq \left(\frac{1}{1+\gamma\mu}\right)^{2t} \|x_0 - x_\star\|^2 + \frac{\gamma\sigma_\star^2}{\gamma\mu^2 + 2\mu}.$$

1674 F.3 MISSING PROOF OF THEOREM 1

1676 We first prove the following useful lemma.

Proof. By assumption,

$$\phi_{\xi}(y) + \langle \nabla \phi_{\xi}(y), x - y \rangle + \frac{\mu_{\xi}}{2} \|x - y\|^2 \le \phi_{\xi}(x), \quad \text{for almost all } \xi \in \mathcal{D}, \forall x, y \in \mathbb{R}^d.$$

This means that

$$\mathbb{E}_{\xi \sim \mathcal{D}}\left[w_{\xi}\left(\phi_{\xi}(y) + \langle \nabla \phi_{\xi}(y), x - y \rangle + \frac{\mu_{\xi}}{2} \|x - y\|^{2}\right)\right] \leq \mathbb{E}_{\xi \sim \mathcal{D}}\left[w_{\xi}\phi_{\xi}(x)\right], \quad \forall x, y \in \mathbb{R}^{d},$$

which is equivalent to

$$\phi(y) + \langle \nabla \phi(y), x - y \rangle + \frac{\mathbb{E}_{\xi \sim \mathcal{D}} \left[w_{\xi} \mu_{\xi} \right]}{2} \| x - y \|^2 \le \phi(x), \quad \forall x, y \in \mathbb{R}^d$$

1693 So, ϕ is μ -strongly convex.

1696 Now, we are ready to prove our main Theorem 1.

Proof. Let C be any (necessarily nonempty) subset of [n] such that $p_C > 0$. Recall that in view of Equation (8) we have

$$f_C(x) = \mathbb{E}_{\xi \sim \mathcal{D}} \left[\frac{I\left(\xi \in C\right)}{p_{\xi}} f_{\xi}(x) \right]$$

i.e., f_C is a conic combination of the functions $\{f_{\xi} : \xi \in C\}$ with weights $w_{\xi} = \frac{I(\xi \in C)}{p_{\xi}}$. Since each f_{ξ} is μ_{ξ} -strongly convex, Appendix F.3 says that f_C is μ_C -strongly convex with

$$\mu_C := \mathbb{E}_{\xi \sim \mathcal{D}} \left[\frac{I\left(\xi \in C\right) \mu_{\xi}}{p_{\xi}} \right].$$

1708 So, every such f_C is μ -strongly convex with

$$\mu = \mu_{\mathrm{AS}} := \min_{C \subseteq [n], p_C > 0} \mathbb{E}_{\xi \sim \mathcal{D}} \left[\frac{I\left(\xi \in C\right) \mu_{\xi}}{p_{\xi}} \right]$$

1713 Further, the quantity σ_{\star}^2 from (2.3) is equal to

$$\sigma_{\star}^{2} := \mathbf{E}_{\xi \sim \mathcal{D}} \left[\left\| \nabla f_{\xi} \left(x_{\star} \right) \right\|^{2} \right] \stackrel{Eqn. (10)}{=} \sum_{C \subseteq [n], p_{C} > 0} p_{C} \left\| \nabla f_{C} \left(x_{\star} \right) \right\|^{2} := \sigma_{\star, \mathrm{AS}}^{2}.$$

1718 Incorporating Appendix F.2 into the above equation, we prove the theorem.

(6)

1720 F.4 THEORY FOR EXPECTATION FORMULATION

We will formally define our optimization objective, focusing on minimization in expectation form.
 We consider

where $f_{\xi} : \mathbb{R}^d \to \mathbb{R}, \xi \sim \mathcal{D}$ is a random variable following distribution \mathcal{D} .

Assumption F.6. Function $f_{\xi} : \mathbb{R}^d \to \mathbb{R}$ is differentiable for almost all samples $\xi \sim \mathcal{D}$.

 $\min_{x \in \mathbb{R}^d} f(x) \coloneqq \mathbb{E}_{\xi \sim \mathcal{D}} \left[f_{\xi}(x) \right],$

This implies that f is differentiable. We will implicitly assume that the order of differentiation and expectation can be swapped ¹, which means that

$$\nabla f(x) \stackrel{Eqn. (1)}{=} \nabla \mathbb{E}_{\xi \sim \mathcal{D}} \left[f_{\xi}(x) \right] = \mathbb{E}_{\xi \sim \mathcal{D}} \left[\nabla f_{\xi}(x) \right]$$

Assumption F.7. Function $f_{\xi} : \mathbb{R}^d \to \mathbb{R}$ is μ -strongly convex for almost all samples $\xi \sim \mathcal{D}$, where $\mu > 0$. That is

 $f_{\xi}(y) + \langle \nabla f_{\xi}, x - y \rangle + \frac{\mu}{2} ||x - y||^2 \le f_{\xi}(x),$

for all $x, y \in \mathbb{R}^d$.

This implies that f is μ -strongly convex, and hence f has a unique minimizer, which we denote by x_{\star} . We know that $\nabla f(x_{\star}) = 0$. Notably, we do *not* assume f to be L-smooth.

Let S be a probability distribution over all *finite* subsets of \mathbb{N} . Given a random set $S \sim S$, we define $p_i \coloneqq \operatorname{Prob}(i \in S), \quad i \in \mathbb{N}.$

We will restrict our attention to proper and nonvacuous random sets.

Assumption F.8. S is proper (i.e., $p_i > 0$ for all $i \in \mathbb{N}$) and nonvacuous (i.e., $\operatorname{Prob}(S = \emptyset) = 0$).

Let C be the selected cohort. Given $\emptyset \neq C \subset \mathbb{N}$ and $i \in \mathbb{N}$, we define

$$v_i(C) := \begin{cases} \frac{1}{p_i} & i \in C\\ 0 & i \notin C, \end{cases}$$
(7)

and

$$f_C(x) := \mathbb{E}_{\xi \sim \mathcal{D}} \left[v_{\xi}(C) f_{\xi}(x) \right] \stackrel{Eqn. (7)}{=} \mathbb{E}_{\xi \sim \mathcal{D}} \left[\frac{I\left(\xi \in C\right)}{p_{\xi}} f_{\xi}(x) \right].$$
(8)

Note that $v_i(S)$ is a random variable and f_S is a random function. By construction, $E_{S \sim S}[v_i(S)] =$ 1 for all $i \in \mathbb{N}$, and hence

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$$\mathbb{E}_{S \sim \mathcal{S}} \left[f_S(x) \right] = \mathbb{E}_{S \sim \mathcal{S}} \left[\mathbb{E}_{\xi \sim \mathcal{D}} \left[v_{\xi}(C) \nabla f_{\xi}(x) \right] \right]$$

$$= \mathbb{E}_{\xi \sim \mathcal{D}} \left[\mathbb{E}_{S \sim \mathcal{S}} \left[v_{\xi}(S) \right] \nabla f_{\xi}(x) \right] = \mathbb{E}_{\xi \sim \mathcal{D}} \left[f_{\xi}(x) \right] = f(x).$$

Therefore, the optimization problem in Equation (1) is equivalent to the stochastic optimization problem

$$\min_{x \in \mathbb{R}^d} \left\{ f(x) := \mathcal{E}_{S \sim \mathcal{S}} \left[f_S(x) \right] \right\}.$$
(9)

Further, if for each $C \subset \mathbb{N}$ we let $p_C := \operatorname{Prob}(S = C)$, f can be written in the equivalent form

$$f(x) = \mathbb{E}_{S \sim \mathcal{S}} \left[f_S(x) \right] = \sum_{C \subset \mathbb{N}} p_C f_C(x) = \sum_{C \subset \mathbb{N}, p_C > 0} p_C f_C(x).$$
(10)

Theorem F.9 (Main Theorem). Let Assumption 2.1 (diferentiability) and Assumption 2.2 (strong convexity) hold. Let S be a random set satisfying Assumption 2.3, and define

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$$\mu_{AS} := \min_{C \subset \mathbb{N}, p_C > 0} \mathbb{E}_{\xi \sim \mathcal{D}} \left[\frac{I(\xi \in C) \, \mu_{\xi}}{p_{\xi}} \right],$$
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$$\mu_{AS} := \sum_{C \subset \mathbb{N}, p_C > 0} p_C \, \|\nabla f_C(x_{\star})\|^2.$$
(11)

Let $x_0 \in \mathbb{R}^d$ be an arbitrary starting point. Then for any $t \ge 0$ and any $\gamma > 0$, the iterates of SPPM-AS (Algorithm 1) satisfy

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1779
$$E\left[\|x_t - x_\star\|^2\right] \le \left(\frac{1}{1 + \gamma\mu_{\rm AS}}\right)^{2t} \|x_0 - x_\star\|^2 + \frac{\gamma\sigma_{\star,\rm AS}^2}{\gamma\mu_{\rm AS}^2 + 2\mu_{\rm AS}}$$

¹This assumption satisfies the conditions required for the theorem about differentiating an integral with a parameter (Fact F.1).

1782 F.5 MISSING PROOF OF ITERATION COMPLEXITY OF SPPM-AS

1784 We have seen above that accuracy arbitrarily close to (but not reaching) $\sigma_{\star,AS}^2/\mu_{AS}^2$ can be achieved 1785 via a single step of the method, provided the stepsize γ is large enough. Assume now that we aim 1786 for ϵ accuracy where $\epsilon \leq \sigma_{\star,AS}^2/\mu_{AS}^2$. Using the inequality $1 - k \leq \exp(-k)$ which holds for all 1787 k > 0, we get

$$\left(\frac{1}{1+\gamma\mu_{\rm AS}}\right)^{2t} = \left(1 - \frac{\gamma\mu}{1+\gamma\mu_{\rm AS}}\right)^{2t} \le \exp\left(-\frac{2\gamma\mu_{\rm AS}t}{1+\gamma\mu_{\rm AS}}\right)^{2t}$$

Therefore, provided that

$$t \ge \frac{1 + \gamma \mu_{\rm AS}}{2\gamma \mu_{\rm AS}} \log \left(\frac{2 \left\| x_0 - x_\star \right\|^2}{\varepsilon} \right)$$

1797 we get $\left(\frac{1}{1+\gamma\mu_{AS}}\right)^{2t} \|x_0 - x_\star\|^2 \le \frac{\varepsilon}{2}$. Furthermore, as long as $\gamma \le \frac{2\varepsilon\mu_{AS}}{2\sigma_{\star,AS}^2 - \varepsilon\mu_{AS}^2}$ (this is true pro-1798 vided that the more restrictive but also more elegant-looking condition $\gamma \le \frac{\varepsilon\mu_{AS}}{\sigma_{\star,AS}^2 - \varepsilon\mu_{AS}^2}$ holds), 1799 we get $\frac{\gamma\sigma_{\star,AS}^2}{\gamma\mu_{AS}^2 + 2\mu_{AS}} \le \frac{\varepsilon}{2}$. Putting these observations together, we conclude that with the stepsize 1801 $\gamma = \frac{\varepsilon\mu_{AS}}{\sigma_{\star,AS}^2}$, we get $\mathbb{E}\left[\|x_t - x_\star\|^2\right] \le \varepsilon$ provided that

$$t \ge \frac{1 + \gamma \mu_{\mathrm{AS}}}{2\gamma \mu_{\mathrm{AS}}} \log \frac{2 \left\| x_0 - x_\star \right\|^2}{\varepsilon} = \left(\frac{\sigma_{\star,\mathrm{AS}}^2}{2\varepsilon \mu_{\mathrm{AS}}^2} + \frac{1}{2} \right) \log \left(\frac{2 \left\| x_0 - x_\star \right\|^2}{\varepsilon} \right)$$

1807 F.6 $\sigma_{\star,\text{NICE}}^2(\tau)$ and $\mu_{\text{NICE}}(\tau)$ are Monotonous Functions of τ

For all $0 \le \tau \le n - 1$:

1.
$$\mu_{\text{NICE}}(\tau+1) \ge \mu_{\text{NICE}}(\tau)$$
,

2.
$$\sigma_{\star,\text{NICE}}^2(\tau) = \frac{\frac{n_{\tau}-1}{n-1}}{n-1} \sigma_{\star,\text{NICE}}^2(1) \le \frac{1}{\tau} \sigma_{\star,\text{NICE}}^2(1).$$

Proof. 1. Pick any $1 \le \tau < n$, and consider a set C for which the minimum is attained in

$$\mu_{\text{NICE}}(\tau+1) = \min_{C \subseteq [n], |C| = \tau+1} \frac{1}{\tau+1} \sum_{i \in C} \mu_i$$

Let $j = \arg \max_{i \in C} \mu_i$. That is, $\mu_j \ge \mu_i$ for all $i \in C$. Let C_j be the set obtained from C by removing the element j. Then $|C_j| = \tau$ and

$$\mu_j = \max_{i \in C} \mu_i \ge \max_{i \in C_j} \mu_i \ge \frac{1}{\tau} \sum_{i \in C_j} \mu_i.$$

By adding $\sum_{i \in C_i} \mu_i$ to the above inequality, we obtain

$$\mu_j + \sum_{i \in C_j} \mu_i \ge \frac{1}{\tau} \sum_{i \in C_j} \mu_i + \sum_{i \in C_j} \mu_i.$$

Observe that the left-hand side is equal to $\sum_{i \in C} \mu_i$, and the right-hand side is equal to $\frac{\tau+1}{\tau} \sum_{i \in C_i} \mu_i$. If we divide both sides by $\tau + 1$, we obtain

$$\frac{1}{\tau+1}\sum_{i\in C}\mu_i \ge \frac{1}{\tau}\sum_{i\in C_j}\mu_i.$$

Since the left-hand side is equal to $\mu_{\text{NICE}}(\tau+1)$, and the right hand side is an upper bound on $\mu_{\text{NICE}}(\tau)$, we conclude that $\mu_{\text{NICE}}(\tau+1) \ge \mu_{\text{NICE}}(\tau)$.

2. In view of (8) we have

 $f_C(x) = \sum_{i \in C} \frac{1}{np_i} f_i(x).$ (12)

$$\sigma_{\star,AS}^{2} = \mathbf{E}_{S\sim\mathcal{S}} \left[\left\| \sum_{i\in S} \frac{1}{np_{i}} \nabla f_{i}(x_{\star}) \right\|^{2} \right] = \mathbf{E}_{S\sim\mathcal{S}} \left[\left\| \sum_{i\in S} \frac{1}{\tau} \nabla f_{i}(x_{\star}) \right\|^{2} \right]$$
(13)

Let χ_i be the random variable defined by

$$\chi_j = \begin{cases} 1 & j \in S \\ 0 & j \notin S. \end{cases}$$
(14)

It is easy to show that

$$\mathbb{E}[\chi_j] = \operatorname{Prob}(j \in S) = \frac{\tau}{n}.$$
(15)

Let fix the cohort S. Let χ_{ij} be the random variable defined by

$$\chi_{ij} = \begin{cases} 1 & i \in S \text{ and } j \in S \\ 0 & \text{otherwise.} \end{cases}$$
(16)

Note that

$$\chi_{ij} = \chi_i \chi_j. \tag{17}$$

Further, it is easy to show that

$$\mathbb{E}[\chi_{ij}] = \operatorname{Prob}(i \in S, j \in S) = \frac{\tau(\tau - 1)}{n(n - 1)}.$$
(18)

Denote $a_i := \nabla f_i(x_\star)$.

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1892		Гц ц2Л		Ги и2]
1893		$\mathbb{T}\left[\left\ \frac{1}{\sum}\right\ ^{2}\right]$		$1_{\mathbb{T}} \left\ \sum_{i=1}^{n} \right\ _{i}^{2}$
1894		$\mathbb{E}\left[\left\ \frac{-}{\tau}\sum_{i\in G}a_{i}\right\ \right]$	=	$\overline{\tau^2}^{\mathbb{E}} \left\ \sum_{i=\sigma}^{\infty} a_i \right\ $
1090				
1090				$1 - \left\ \sum_{n=1}^{n} \right\ ^2$
1097			=	$\frac{1}{\tau^2}\mathbb{E}\left[\left\ \sum \chi_i a_i\right\ \right]$
1890				, $\lfloor i=1$ $ \rfloor$
1900				$1 \begin{bmatrix} n \end{bmatrix}$
1901			=	$\frac{1}{2}\mathbb{E}\left \sum \ \chi_{i}a_{i}\ ^{2} + \sum \langle \chi_{i}a_{i}, \chi_{j}a_{j} \rangle\right $
1902				$7^2 \begin{bmatrix} 1 & 1 & 1 \\ i=1 & i\neq j \end{bmatrix}$
1903				$\begin{bmatrix} n \\ \end{bmatrix}$
1904			=	$\frac{1}{2} \mathbb{E} \left[\sum_{i=1}^{n} \ \chi_{i} a_{i} \ ^{2} + \sum_{i=1}^{n} \chi_{ii} \langle a_{i}, a_{i} \rangle \right]$
1905				$\tau^{2} \left[\sum_{i=1}^{2} \frac{\chi_{i} \chi_{i} \chi_{i} \chi_{i}}{i \neq j} \frac{\chi_{i} \chi_{i} \chi_{i} \chi_{i} \chi_{i}}{i \neq j} \right]$
1906				
1907			=	$\frac{1}{2}\sum \mathbb{E}[\chi_i] \ a_i\ ^2 + \sum \mathbb{E}[\chi_{ij}] \langle a_i, a_j \rangle$
1908				$\tau^2 \sum_{i=1}^{i=1} \sum_{j\neq j} \sum_{i\neq j} \sum_{j\neq j} \sum$
1909				$\begin{pmatrix} n \end{pmatrix}$
1910			=	$\frac{1}{\tau} \left(\frac{\tau}{\tau} \sum_{i=1}^{n} a_i ^2 + \frac{\tau(\tau-1)}{\tau} \sum_{i=1}^{n} \langle a_{ii}, a_{ii} \rangle \right)$
1911				$\tau^2 \left(n \sum_{i=1}^{n} n(n-1) \sum_{i\neq j} (\omega_i, \omega_j) \right)$
1912				
1913			=	$\frac{1}{2}\sum \ a_i\ ^2 + \frac{\tau-1}{2}\sum \langle a_i, a_i \rangle$
1914				$\tau n \sum_{i=1}^{n} \tau n(n-1) \sum_{i \neq j} \tau n(n-1)$
1915				$n \qquad (\parallel n \parallel \parallel^2 n)$
1916			=	$\frac{1}{1} \sum \ a_i\ ^2 + \frac{\tau - 1}{1} \left(\ \sum a_i\ - \sum \ a_i\ ^2 \right)$
1917				$\tau n \sum_{i=1}^{n} \tau n(n-1) \left(\left\ \sum_{i=1}^{n} \right\ \sum_{i=1}^{n} \left\ \sum_{i=1}^{n} \right\ \right)$
1918				n (n $ _{2}$
1919			=	$\frac{n-\tau}{(n-\tau)} \frac{1}{2} \sum a_i ^2 + \frac{n(\tau-1)}{(n-\tau)} \left\ \frac{1}{2} \sum a_i \right\ $
1920				$\tau(n-1) n \underset{i=1}{\overset{n}{\rightharpoonup}} \tau(n-1) \left\ n \underset{i=1}{\overset{n}{\rightharpoonup}} \tau \right\ $
1921				$m = 1^{n}$ $m(= 1) \parallel 1^{n}$ \parallel^2
1923			=	$\frac{n-\tau}{(1-\tau)} \frac{1}{\tau} \sum_{i} \ \nabla f_i(x_{\star})\ ^2 + \frac{n(\tau-1)}{(1-\tau)} \ \frac{1}{\tau} \sum_{i} \nabla f_i(x_{\star})\ $
1924				$\tau(n-1) n \sum_{i=1}^{n} \tau(n-1) \ n \sum_{i=1}^{n} \tau(n-1) \ \ n \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ $
1925				$n-\tau$ 1 $\sum_{n=1}^{n}$ (τ) (τ)
1926			=	$\frac{1}{\tau(n-1)} \frac{1}{n} \sum \ \nabla f_i(x_\star)\ ^2$
1927				i=1
1928			<	$\frac{1}{2} \frac{1}{2} \sum_{n=1}^{n} \ \nabla f_{n}(x_{n})\ ^{2}$
1929			_	$\tau n \sum_{i=1}^{n} \nabla J_i(w_\star) $
1930				0 —1
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1936	г./	WIISSING PROOF OF LI	1MMA	A 1
1937				

For ease of notation, let $a_i = \nabla f_i(x_{\star})$ and $\hat{z}_j = |C_j| a_{\xi_j}$, and recall that 1938 1939

 $\sigma_{\star,\mathrm{SS}}^2 = \mathrm{E}_{\xi_1,\ldots,\xi_b} \left[\left\| \frac{1}{n} \sum_{j=1}^b \hat{z}_j \right\|^2 \right].$ 41 1942 (19) 1943

where $\xi_j \in C_j$ is chosen uniformly at random. Further, for each $j \in [b]$, let $z_j = \sum_{i \in C_j} a_i$. Observe 1945 that $\sum_{j=1}^{b} z_j = \sum_{j=1}^{b} \sum_{i \in C_j} a_i = \sum_{i=1}^{n} a_i = \nabla f(x_*) = 0$. Therefore, 1946 1947 $\left\|\frac{1}{n}\sum_{i=1}^{b}\hat{z}_{i}\right\|^{2} = \frac{1}{n^{2}}\left\|\sum_{j=1}^{b}\hat{z}_{j} - \sum_{j=1}^{b}z_{j}\right\|^{2}$ 1948 1949 1950 $= \frac{b^2}{n^2} \left\| \frac{1}{b} \sum_{i=1}^{b} \left(\hat{z}_j - z_j \right) \right\|^2$ 1951 1952 1953 $\leq \frac{b^2}{n^2} \frac{1}{b} \sum_{i=1}^{b} \|\hat{z}_j - z_j\|^2$ 1957 $= \frac{b}{n^2} \sum_{i=1}^{b} \|\hat{z}_j - z_j\|^2,$ (20)1959 where the inequality follows from convexity of the function $u \mapsto ||u||^2$. Next, 1961 1962 $\|\hat{z}_j - z_j\|^2 = \left\| |C_j| \, a_{\xi_j} - \sum_{i \in C_j} a_i \right\|^2 = |C_j|^2 \left\| a_{\xi_j} - \frac{1}{|C_j|} \sum_{i \in C_j} a_i \right\|^2 \le |C_j|^2 \, \sigma_j^2.$ 1963 (21)1964 1965 1966 By combining Equation (19), Equation (20) and Equation (21), we get 1967 $\sigma_{\star,\mathrm{SS}}^2 \stackrel{Eqn. (19)}{=} \mathrm{E}_{\xi_1, \dots, \xi_b} \left[\left\| \frac{1}{n} \sum_{i=1}^b \hat{z}_i \right\|^2 \right]$ 1968 1969 1970 $\stackrel{Eqn. (20)}{\leq} \quad \mathbf{E}_{\xi_{1},...,\xi_{b}} \left[\frac{b}{n^{2}} \sum_{i=1}^{b} \|\hat{z}_{j} - z_{j}\|^{2} \right]$ 1972 1974 $\stackrel{Eqn. (21)}{\leq} \quad \operatorname{E}_{\xi_1, \dots, \xi_b} \left[\frac{b}{n^2} \sum_{i=1}^b \left| C_j \right|^2 \sigma_j^2 \right]$ 1975 1976

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1980
$$= \frac{b}{n^2} \sum_{j=1}^{b} |C_j|^2 \sigma_j^2.$$

1981 1982

1984 1985

1987

1988 1989

1944

The last expression can be further bounded as follows:

$$\frac{b}{n^2} \sum_{j=1}^b |C_j|^2 \, \sigma_j^2 \le \frac{b}{n^2} \left(\sum_{j=1}^b |C_j|^2 \right) \max_j \sigma_j^2 \le \frac{b}{n^2} \left(\sum_{j=1}^b |C_j| \right)^2 \max_j \sigma_j^2 = b \max_j \sigma_j^2,$$

where the second inequality follows from the relation $||u||_2 \leq ||u||_1$ between the L_2 and L_1 norms, and the last identity follows from the fact that $\sum_{i=1}^{b} |C_i| = n$.

F.8 STRATIFIED SAMPLING AGAINST BLOCK SAMPLING AND NICE SAMPLING 1990

In this section, we present a theoretical comparison of block sampling and its counterparts, providing 1992 a theoretical justification for selecting block sampling as the default clustering method in future 1993 experiments. Additionally, we compare various sampling methods, all with the same sampling size, b: b-nice sampling, block sampling with b clusters, and block sampling, where all clusters are of uniform size b. 1996

Assumption F.10. For simplicity of comparison, we assume b clusters, each of the same size, b: 1997

$$|C_1| = |C_2| = \ldots = |C_b| = b$$

It is crucial to acknowledge that, without specific assumptions, the comparison of different sampling methods may not provide meaningful insights. For instance, the scenario described in Lemma 1, characterized by complete inter-cluster homogeneity, demonstrates that stratified sampling achieves a variance term, denoted as $\sigma_{\star,SS}^2$, which is lower than the variance terms associated with both block sampling and nice sampling. However, a subsequent example illustrates examples in which the variance term for block sampling surpasses those of block sampling and nice sampling. Without imposing any additional clustering assumptions, there exist examples for any arbitrary n, such that $\sigma_{\star,\mathrm{SS}}^2 \geq \sigma_{\star,\mathrm{BS}}^2$ and $\sigma_{\star,\mathrm{SS}}^2 \geq \sigma_{\star,\mathrm{NICE}}^2$.

Proof. Counterexample when SS is worse in neighborhood than BS

Assume we have such clustering and $\nabla f_i(x_*)$ such that the centroids of each cluster are equal to zero: $\forall i \in [b], \frac{1}{|C_i|} \sum_{j \in C_i} \nabla f_j(x_\star) = 0$. For instance, this can be achieved in the following case: The dimension is d = 2, all clusters are of equal size m, then assign $\forall i \in [b], \forall j \in C_i$, $\nabla f_i(x_\star) = \left(Re\left(\omega^{mj+i}\right), Im\left(\omega^{mj+i}\right) \right)$ where $\omega = \sqrt[n]{1} \in \mathbb{C}$. Let us calculate $\sigma^2_{\star, BS}$:

$$\sigma_{\star,\mathrm{BS}}^2 := \sum_{j=1}^b q_j \left\| \sum_{i \in C_j} \frac{1}{np_i} \nabla f_i(x_\star) \right\|^2 =$$

$$= \frac{1}{n^2} \sum_{j=1}^{b} \frac{|C_j|^2}{q_j} \left\| \frac{1}{|C_j|} \sum_{i \in C_j} \nabla f_i(x_\star) \right\|^2 = 0.$$

 $\sigma_{+BS}^2 = 0 < \sigma_{+SS}^2$.

As a result:

Counterexample when SS is worse in neighborhood than NICE
Here, we employ a similar proof technique as in the proof of Lemma 2. Let us choose such clustering
$$C_{b,SS,max} = \arg \max_{C_b} \sigma_{\star,SS}^2(C_b)$$
. Denote $\mathbf{i}_b := (i_1, \cdots, i_b)$, $\mathbf{C}_b := C_1 \times \cdots \times C_b$, and $S_{\mathbf{i}_b} :=$
 $\|\frac{1}{\tau} \sum_{i \in \mathbf{i}_b} \nabla f_i(x_\star)\|$.

$$\stackrel{1}{=} \frac{1}{\#_{\text{clusterizations}}} \sum_{C_b} \frac{1}{b^b} \sum_{\mathbf{i}_b \in \mathbf{C}_b} S_{\mathbf{i}_b}$$

$$= \frac{1}{\mu} \sum \sigma_{\star,\mathrm{SS}}^2(\mathcal{C}_b)$$

 $=\frac{1}{C(n,b)}\sum_{\mathbf{i}_b\subset[n]}S_{\mathbf{i}_b}$

$$= \frac{1}{\#_{\text{clusterizations}}} \sum_{C_b} \sigma_{\star,i}^2$$

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$$\stackrel{2}{\leq} \sigma^2_{\star,\mathrm{SS}}(\mathcal{C}_{b,\mathrm{SS},\mathrm{max}}).$$

Equation 1 holds because, in every clusterization C_b , there are $\frac{1}{b^b}$ possible sample combinations \mathbf{i}_b . Due to symmetry, one can conclude that each combination $S_{\mathbf{i}_b}$ is counted the same number of times. Equation 2 follows from the definition of $C_{b,SS,max}$.

For illustrative purposes, we can demonstrate this effect with a specific example. Let n = 4 and define $\forall i \ a_i = \nabla f_i(x^*) \in \mathbb{R}^2$. Let $a_1 = (0, 1)^T$, $a_2 = (1, 0)^T$, $a_3 = (0, -1)^T$, and $a_4 = (-1, 0)^T$. Then fix clustering $C_b = \{C_1 = \{a_1, a_3\}, C_2 = \{a_2, a_4\}\}$. Then:

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$$\sigma_{\star,SS}^2 = \frac{1}{4} \sum_{\mathbf{i}_b \in \mathcal{C}_b} \left\| \frac{a_{i_1} + a_{i_2}}{2} \right\|^2$$

2048
2049
$$= \frac{1}{4} \sum \left\| (\pm \frac{1}{2}, \pm \frac{1}{2}) \right\|^2$$

2049
$$= -\frac{1}{4} \sum_{\mathbf{i}_b \in \mathcal{C}_b} \left\| (\pm \frac{1}{2}, \pm \frac{1}{2}) \right\|$$

$$\begin{array}{l} 2050 \\ 2051 \end{array} = \frac{1}{2}. \end{array}$$

2052	$1 - \ a_i + a_j\ ^2$
2053	$\sigma_{\star,\text{NICE}}^2 = \frac{1}{C(4,2)} \sum \left\ \frac{a_i + a_j}{2} \right\ $
2054	$(1,2)_{i < j} \parallel 2 \parallel$
2055	$1 \sum \ a_i + a_j\ ^2$
2056	$=\frac{1}{6}\sum_{i}\left\ \frac{1}{2}\right\ $
2057	$i < j \parallel$
2058	$1\left(\left\ a_{1}+a_{3}\ ^{2} \ a_{2}+a_{4}\ ^{2}\right) \ a_{i_{1}}+a_{i_{2}}\ ^{2}\right)$
2059	$= \frac{1}{6} \left(\left\ \frac{1}{2} + \frac{1}{2} \right\ + \left\ \frac{2}{2} + \frac{1}{2} \right\ + 2 \times \left\ \frac{1}{2} + \frac{1}{2} \right\ \right)$
2060	
2061	$-\frac{1}{2}\left(0+2\times2\times\frac{1}{2}\right)$
2062	$= \frac{6}{6} \left(\frac{6+2\times2\times2}{2} \right)$
2063	1
2064	$=\frac{1}{3}$
2065	$\frac{1}{2}$
2066	$= \frac{1}{3} \times \sigma_{\star, SS}^2$
2067	$< \sigma^2$ reg
2068	— * *,55
2069	
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2073	\mathbf{T} , \mathbf
2074	To select the optimal clustering, we will choose the clustering that minimizes $\sigma_{\star,SS}^{*}$.
2076 2077 2078	blocks as $C_b := \{C_1, C_2, \dots, C_b\}$, such that the disjoint union of all clusters $C_1 \cup C_2 \cup \dots \cup C_b = [n]$. Define <i>block sampling Optimal Clustering</i> as the clustering configuration that minimizes $\sigma^2_{\star,SS}$, formally given by:
2079	$\mathcal{C}_{h \text{ SS}} := \arg\min \sigma_{+\text{ SS}}^2(\mathcal{C}_h).$
2000	\mathcal{C}_b
2001	
2002	Proof of the Lemma 2 is provided below
2003	1 tool of the Lemma 2 is provided below.
2004	
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2087	<i>Proof.</i> 1. Denote $\mathbf{i}_b \coloneqq (i_1, \dots, i_b)$, $\mathbf{C}_b \coloneqq C_1 \times \dots \times C_b$, and $S_{\mathbf{i}_b} \coloneqq \left\ \frac{1}{\tau} \sum_{i \in \mathbf{i}} \nabla f_i(x_*) \right\ $.
2088	
2089	$ \cdot ^2$
2090	$\sigma^2_{\text{NGD}} = \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} \left\ \frac{1}{2} \sum_{i=1}^{n} \nabla f_i(x_i) \right\ $
2091	$C(n,\tau) \simeq C(n,\tau) \simeq C(n) \left\ \tau \simeq C(n,\tau) \right\ _{C \subseteq [n]} \left\ \tau \simeq C(n,\tau) \right\ _{C \subseteq [n]}$
2092	
2093	$=\frac{1}{C(n-b)}\sum S_{\mathbf{i}_b}$
2094	$\bigcup (n, b) \frac{1}{\mathbf{i}_b \subseteq [n]}$
2095	$1 \qquad 1 \qquad \sum 1 \sum c$
2096	$=\frac{1}{\#_{\text{clusterizations}}}\sum_{a}\frac{1}{b^{b}}\sum_{a}S_{\mathbf{i}_{b}}$
2097	\mathcal{C}_b $\mathbf{i}_b \in \mathbf{C}_b$
2098	$=\frac{1}{\mu}\sum_{a}\sigma_{+SS}^{2}(\mathcal{C}_{b})$
2099	#clusterizations \mathcal{L}_{b}

2100 $\geq \sigma_{\star,\mathrm{SS}}^2(\mathcal{C}_{b,\mathrm{SS},\mathrm{min}})$

2102

2103 Equation 1 holds because, in every clusterization C_b , there are $\frac{1}{b^b}$ possible sample combi-2104 nations \mathbf{i}_b . Due to symmetry, one can conclude that each combination $S_{\mathbf{i}_b}$ is counted the 2105 same number of times. Equation 2 follows from the definition of $C_{b,SS,min}$ as the clustering that minimizes $\sigma_{\star,SS}^2$, according to Definition F.11. $\mu_{\text{NICE}(b)} := \min_{\substack{C \subseteq [n] \\ |C| = b}} \frac{1}{b} \sum_{i \in C} \mu_i,$

2. The neighborhood size for SPPM-AS is given by $\frac{\gamma \sigma_{\star,AS}^2}{\gamma \mu_{AS}^2 + 2\mu_{AS}}$, denoted as U_{AS} for simplicity. Define:

Using the definition of the set $\mathbf{C}_b := C_1 \times C_2 \times \cdots \times C_b$, we have $\mathbf{C}_b \subseteq \{C \subseteq [n] \mid |C| =$ *b*}. Applying this fact, we obtain:

 $\mu_{\rm SS} := \min_{\mathbf{i}_b \in \mathbf{C}_b} \sum_{i_j=1}^b \frac{\mu_{i_j} |C_j|}{n} \stackrel{\text{Asm. 10}}{=} \min_{\mathbf{i}_b \in \mathbf{C}_b} \sum_{i_j=1}^b \frac{\mu_{i_j} b}{b^2} = \min_{\mathbf{i}_b \in \mathbf{C}_b} \frac{1}{b} \sum_{i_j=1}^b \mu_{i_j}.$

$$\mu_{\rm SS} = \min_{\mathbf{i}_b \in \mathbf{C}_b} \frac{1}{b} \sum_{j \in \mathbf{i}_b} \mu_j \ge \mu_{\rm NICE(b)}$$

Combining the above with $\sigma^2_{\star,SS}(\mathcal{C}_{b,SS}) \leq \sigma^2_{\star,NICE}$, we obtain that $U_{SS}(\mathcal{C}_{b,SS}) \leq U_{NICE}$, demonstrating the variance reduction of SS compared to NICE.

Consider the number of clusters and the size of each cluster, with b = 2, under Assumption F.10. Then, $\sigma_{\star,\mathrm{SS}}^2(\mathcal{C}_{b,\mathrm{SS}}) \leq \sigma_{\star,\mathrm{BS}}^2$.

 $\sigma_{\star,\mathrm{SS}}^2 = \frac{1}{4} \left(S^2 - \left\| \frac{a_{C_1^1} + a_{C_1^2}}{2} \right\|^2 - \left\| \frac{a_{C_2^1} + a_{C_2^2}}{2} \right\|^2 \right)$

Proof. Let n = 4, b = 2. Denote $\forall i \ a_i = \nabla f_i(x_*)$. Define $S^2 := \sum_{i < j} \left\| \frac{a_i + a_j}{2} \right\|^2$.

 $C_{b,SS}$ clustering minimizes $\sigma_{\star,SS}^2$, thereby maximizing $\sigma_{\star,BS}^2$. Thus,

 $=\frac{1}{4}\left(S^2-2\sigma_{\star,\mathrm{BS}}^2\right)$

$$\begin{split} &\sigma_{\star,\mathrm{SS}}^2 = \frac{1}{4} \left(\left[\left\| \frac{a_{C_1^1} + a_{C_2^1}}{2} \right\|^2 + \left\| \frac{a_{C_1^2} + a_{C_2^2}}{2} \right\|^2 \right] + \left[\left\| \frac{a_{C_1^1} + a_{C_2^2}}{2} \right\|^2 + \left\| \frac{a_{C_1^2} + a_{C_2^1}}{2} \right\|^2 \right] \right) \\ &= \frac{1}{4} \left(2\sigma_{\star,\mathrm{BS}}^2 \left((C_1^1, C_2^1), (C_1^2, C_2^2) \right) + 2\sigma_{\star,\mathrm{BS}}^2 \left((C_1^1, C_2^2), (C_1^2, C_2^1) \right) \right) \\ &= \frac{1}{2} \left(\sigma_{\star,\mathrm{BS}}^2 \left((C_1^1, C_2^1), (C_1^2, C_2^2) \right) + \sigma_{\star,\mathrm{BS}}^2 \left((C_1^1, C_2^2), (C_1^2, C_2^1) \right) \right) \\ &\leq \sigma_{\star,\mathrm{BS}}^2. \end{split}$$

However, it is possible that this relationship might hold more generally. Empirical experiments for different configurations, such as b = 3, support this possibility. For example, with n = 9, b = 3, and d = 10, Python simulations where gradients ∇f_i are sampled from $\mathcal{N}(0,1)$ and $\mathcal{N}(e,1)$ across 1000 independent trials, show that $\sigma^2_{\star,SS} \leq \sigma^2_{\star,BS}$. Question of finding theoretical proof for arbitrary *n* remains open and has yet to be addressed in the existing literature.

F.9 DIFFERENT APPROACHES OF FEDERATED AVERAGING

Proof of Theorem 2:

Proof. $\|x_t - x_\star\|^2 = \left\|\sum_{i \in S} \frac{1}{|S_t|} \operatorname{prox}_{\gamma f_i}(x_{t-1}) - \frac{1}{|S_t|} \sum_{i \in S} x_\star\right\|^2$ $\stackrel{(Fact \ F.3)}{=} \left\| \sum_{i \in S_{t}} \frac{1}{|S_{t}|} \left[\operatorname{prox}_{\gamma f_{i}}(x_{t-1}) - \operatorname{prox}_{\gamma f_{i}}(x_{\star} + \gamma \nabla f_{i}(x_{\star})) \right] \right\|^{2}$ $\overset{Jensen}{\leq} \sum_{i \in S_t} \frac{1}{|S_t|} \left\| \left[\operatorname{prox}_{\gamma f_i}(x_{t-1}) - \operatorname{prox}_{\gamma f_i}(x_\star + \gamma \nabla f_i(x_\star)) \right] \right\|^2$ $\leq \sum_{i \in S_{\star}} \frac{1}{|S_{t}|} \frac{1}{(1+\gamma\mu_{i})^{2}} \|x_{t-1} - (x_{\star} + \gamma\nabla f_{i}(x_{\star}))\|^{2}$

$$\mathbb{E}_{S_{t} \sim S} \left[\|x_{t} - x_{\star}\|^{2} |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \frac{1}{(1 + \gamma \mu_{i})^{2}} \|(x_{t-1} - x_{\star}) - \gamma \nabla f_{i}(x_{\star}))\|^{2} |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \frac{1}{(1 + \gamma \mu_{i})^{2}} \left((1 + \alpha_{i}) \|x_{t-1} - x_{\star}\|^{2} + (1 + \alpha_{i}^{-1}) \|\gamma \nabla f_{i}(x_{\star}))\|^{2} \right) |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \frac{1}{(1 + \gamma \mu_{i})^{2}} \left((1 + \gamma \mu_{i}) \|x_{t-1} - x_{\star}\|^{2} + \left(1 + \frac{1}{\gamma \mu_{i}}\right) \|\gamma \nabla f_{i}(x_{\star}))\|^{2} \right) |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \frac{1}{(1 + \gamma \mu_{i})^{2}} \left((1 + \gamma \mu_{i}) \|x_{t-1} - x_{\star}\|^{2} + \left(1 + \frac{1}{\gamma \mu_{i}}\right) \|\gamma \nabla f_{i}(x_{\star}))\|^{2} \right) |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \left(\frac{1}{(1 + \gamma \mu_{i})} \|x_{t-1} - x_{\star}\|^{2} + \frac{\gamma}{(1 + \gamma \mu_{i})\mu_{i}} \|\nabla f_{i}(x_{\star})\|^{2} \right) |x_{t-1} \right]$$

$$\mathbb{E}_{S_{t} \sim S} \left[\sum_{i \in S_{t}} \frac{1}{|S_{t}|} \sum_{i \in S_{t}} \frac{1}{1 + \gamma \mu_{i}} |x_{t-1}|^{2} + \mathbb{E}_{S_{t} \sim S} \left[\frac{1}{|S_{t}|} \sum_{i \in S_{t}} \frac{\gamma}{(1 + \gamma \mu_{i})\mu_{i}} \|\nabla f_{i}(x_{\star})\|^{2} |x_{t-1}|^{2} \right]$$

By applying tower property one can get the following:

$$\begin{split} \mathbb{E}_{S_t \sim \mathcal{S}} \left[\left\| x_t - x_\star \right\|^2 \right] \\ &= \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{1}{1 + \gamma \mu_i} \right] \left\| x_{t-1} - x_\star \right\|^2 + \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{\gamma}{(1 + \gamma \mu_i)\mu_i} \left\| \nabla f_i(x_\star) \right) \right\|^2 \right] \\ &= A_{\mathcal{S}} \left\| x_{t-1} - x_\star \right\|^2 + B_{\mathcal{S}}. \\ &\text{where } A_{\mathcal{S}} \coloneqq \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{1}{1 + \gamma \mu_i} \right] \text{ and } B_{\mathcal{S}} \coloneqq \mathbb{E}_{S_t \sim \mathcal{S}} \left[\frac{1}{|S_t|} \sum_{i \in S_t} \frac{\gamma}{(1 + \gamma \mu_i)\mu_i} \left\| \nabla f_i(x_\star) \right) \right\|^2 \right]. \\ &\text{By directly applying Fact F.5:} \\ & \mathbb{E}_{S_t \sim \mathcal{S}} \left[\left\| x_t - x_\star \right\|^2 \right] \leq A_{\mathcal{S}}^t \| x_0 - x_\star \|^2 + \frac{B_{\mathcal{S}}}{1 - A_{\mathcal{S}}}. \end{split}$$

[Inexact formulation of SPPM-AS] Let $b > 0 \in \mathbb{R}$ and define $\widetilde{\operatorname{prox}}_{\gamma f}(x)$ such that $\forall x \| \widetilde{\text{prox}}_{\gamma f}(x) - \operatorname{prox}_{\gamma f}(x) \|^2 \leq b$. Let Assumption 2.1 and Assumption 2.2 hold. Let $x_0 \in \mathbb{R}^d$ be an arbitrary starting point. Then for any $t \geq 0$ and any $\gamma > 0$, s > 0, the iterates of SPPM-AS satisfy

$$\mathbb{E}\left[\|x_t - x_\star\|^2\right] \le \left(\frac{1+s}{(1+\gamma\mu)^2}\right)^t \|x_0 - x_\star\|^2 + \frac{(1+s)\left(\gamma^2\sigma_\star^2 + s^{-1}b(1+\gamma\mu)^2\right)}{\gamma^2\mu^2 + 2\gamma\mu - s}$$

Proof of Lemma F.9. We provide more general version of SPPM proof

$$\begin{aligned} & 2216 \\ 2217 \\ & \|x_{t+1} - x_{\star}\|^{2} = \left\|\widetilde{\operatorname{prox}}_{\gamma f_{\xi_{t}}(x_{t})} - \operatorname{prox}_{\gamma f_{\xi_{t}}}(x_{t}) + \operatorname{prox}_{\gamma f_{\xi_{t}}}(x_{t}) - x_{\star}\right\|^{2} \\ & 2218 \\ & \leq (1 + s^{-1}) \left\|\widetilde{\operatorname{prox}}_{\gamma f_{\xi_{t}}}(x_{t}) - \operatorname{prox}_{\gamma f_{\xi_{t}}}\right\|^{2} (x_{t}) + (1 + s) \left\|\operatorname{prox}_{\gamma f_{\xi_{t}}}(x_{t}) - x_{\star}\right\|^{2} \\ & 2220 \\ & \leq (1 + s^{-1})b + (1 + s) \left\|\operatorname{prox}_{\gamma f_{\xi_{t}}}(x_{t}) - x_{\star}\right\|^{2}. \end{aligned}$$

Then proof follows same path as proof Theorem 1 and we get

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$$\mathbb{E}\left[\left\|x_{t+1} - x_{\star}\right\|^{2}\right] \leq (1+s^{-1})b + (1+s)\frac{1}{(1+\gamma\mu)^{2}}\left(\left\|x_{t} - x_{\star}\right\|^{2} + \gamma^{2}\sigma_{\star}^{2}\right)$$
$$= \frac{1+s}{(1+\gamma\mu)^{2}}\left(\left\|x_{t} - x_{\star}\right\|^{2} + \left[\gamma^{2}\sigma_{\star}^{2} + s^{-1}b(1+\gamma\mu)^{2}\right]\right).$$

azc It only remains to solve the above recursion. Luckily, that is exactly what Fact F.5 does. In particular, we use it with $s_t = \mathbb{E}\left[\|x_t - x_\star\|^2 \right]$, $A = \frac{1+s}{(1+\gamma\mu)^2}$ and $B = \frac{(1+s)\left(\gamma^2 \sigma_\star^2 + s^{-1}b(1+\gamma\mu)^2\right)}{(1+\gamma\mu)^2}$ to get

> $\mathbb{E}\left[\|x_t - x_\star\|^2\right] \le A^t \|x_0 - x_\star\|^2 + B\frac{1}{1 - A}$ $\leq A^{t} \|x_{0} - x_{\star}\|^{2} + B \frac{(1 + \gamma \mu)^{2}}{(1 + \gamma \mu)^{2} - 1 - s}$ $\leq A^{t} \|x_{0} - x_{\star}\|^{2} + \frac{(1+s)\left(\gamma^{2}\sigma_{\star}^{2} + s^{-1}b(1+\gamma\mu)^{2}\right)}{(1+\gamma\mu)^{2} - 1 - s}$ $= \left(\frac{1+s}{(1+\gamma\mu)^2}\right)^t \|x_0 - x_\star\|^2 + \frac{(1+s)\left(\gamma^2\sigma_\star^2 + s^{-1}b(1+\gamma\mu)^2\right)}{\gamma^2\mu^2 + 2\gamma\mu - s}.$