Label Noise: Ignorance Is Bliss [6 pages]

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Abstract

1	We establish a new theoretical framework for learning under multi-class, instance-
2	dependent label noise. At the heart of our framework is the concept of <i>relative</i>
3	signal strength (RSS), which is a point-wise measure of noisiness. We use relative
4	signal strength to establish matching upper and lower bounds for excess risk . Our
5	theoretical findings reveal a surprising result: the extremely simple Noise Ignorant
6	Empirical Risk Minimization (NI-ERM) principle, which conducts empirical risk
7	minimization as if no label noise exists, is minimax optimal. Finally, we translate
8	these theoretical insights into practice: by using NI-ERM to fit a linear classifier
9	on top of a frozen foundation model, we achieve state-of-the-art performance on
10	the CIFAR-N data challenge.

11 **1 Introduction**

The problem of classification with label noise can be stated in terms of variables (X, Y, \tilde{Y}) , where X is the feature vector, $Y \in \{1, ..., K\}$ is the true label associated to X, and $\tilde{Y} \in \{1, ..., K\}$ as a noisy version of Y. The learner has access to i.i.d. realizations of (X, \tilde{Y}) , and the objective is to learn a classifier that optimizes the risk associated with (X, Y).

In recent years, there has been a surge of interest in the challenging setting of instance (i.e., feature) dependent label noise, in which \tilde{Y} can depend on both Y and X. While several algorithms have been developed, there remains relatively little theory regarding algorithm performance and the fundamental limits of this learning paradigm.

This work develops a theoretical framework for learning under multi-class, instance-dependent label noise. Our framework hinges on the concept of *relative signal strength*, which is a point-wise measure of "noisiness" in a label noise problem. Using relative signal strength, we establish matching upper and lower bounds for excess risk. We further identify distributional assumptions that ensure that the lower bound tends to zero as the sample size *n* grows, implying that consistent learning is possible.

25 Suprisingly, our theoretical findings reveal that Noise Ignorant Empirical Risk Minimization (NI-

ERM), which conducts empirical risk minimization as if no label noise exists, is minimax optimal. To translate this insight into practice, we use NI-ERM to fit a linear classifier on top of a self-supervised

translate this insight into practice, we use NI-ERM to fit a linear classifier on top of a self-sup
 feature extractor, achieving state-of-the-art performance on the CIFAR-N data challenge.

29 2 Literature review

30 Theory and algorithms for classification with label noise are often based on different probabilistic

- models. These can be categorized depending on how \tilde{Y} depends on Y and X. The simplest model is
- symmetric noise, where the distribution of Y is independent of Y and X [Angluin and Laird, 1988].
- In this case, the probability that $\tilde{Y} = k$ is the same for all $k \neq Y$, regardless of Y and X. In this

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setting, it is easy to show that minimizing the noisy excess risk (associated to the 0/1 loss) implies 34

minimizing the clean excess risk, a property known as *immunity*. When immunity holds, there is no 35 need to modify the learning algorithm on account of noisy labels. In other words, the learner may be 36

ignorant of the label noise and still learn consistently. 37

A more general model is classification with label dependent noise, in which the distribution of Y38 depends on Y, but not X. Many practical algorithms have been developed over the years, based 39 40 on principles including data re-weighting [Liu and Tao, 2015], robust training [Han et al., 2018, Liu et al., 2020, Foret et al., 2021] and data cleaning [Brodley and Friedl, 1999, Northcutt et al., 41 2021]. Consistent learning algorithms still exist, such as those based on loss correction [Natarajan 42 et al., 2013, Patrini et al., 2017, Van Rooyen and Williamson, 2018, Liu and Guo, 2020, Zhang et al., 43 2022]. However, the primary limitation of these methods is their reliance on the knowledge of noise 44 transition probabilities, which are generally not identifiable [Zhang et al., 2021b]. 45

In the most general setting, that of instance dependent label noise, the distribution of \tilde{Y} depends on 46 both Y and X. While algorithms are emerging [Cheng et al., 2021, Zhu et al., 2021, Wang et al., 2022, 47 Yang et al., 2023], the theoretical developments have primarily focused on the binary setting. Scott 48 [2019] establishes immunity for a Neyman-Pearson-like performance criterion under a *posterior drift* 49 model, discussed in more detail below. Cannings et al. [2020] establish an upper bound for excess 50 risk under the strong assumption that the optimal classifiers for the clean and noisy distributions are 51 the same. Closest to our work, Im and Grigas [2023] derive excess risk upper and lower bounds, and 52 reach a similar conclusion, that noise-ignorant ERM attains the lower bound. Our results, based on 53 the new concept of relative signal strength, provide a more refined analysis. 54

Additional connections between our contributions and prior work are made throughout the paper. 55

Problem statement and relative signal strength 3 56

Notation. \mathcal{X} denotes a feature space and $\mathcal{Y} = \{1, 2, \dots, K\}$ a label space, with $K \in \mathbb{N}$. The 57 *K*-simplex is $\Delta^K := \{p \in \mathbb{R}^K : \forall i, p_i \ge 0, \sum p_i = 1\}$. A $K \times K$ matrix is *row stochastic* if all of its rows are in Δ^K . Denote the *i*-th element of a vector \boldsymbol{v} as $[\boldsymbol{v}]_i$, and the (i, j)-th element of a 58 59 matrix M as $[M]_{i,j}$. 60

3.1 Learning from label noise 61

In conventional multiclass classification, we observe training data $(X_1, Y_1), \ldots, (X_n, Y_n)$ drawn 62 i.i.d. from a joint distribution P_{XY} . The marginal distribution of X is denoted by P_X , and the class 63 posterior probabilities $P_{Y|X=x}$ are captured by a K-simplex-valued vector $\boldsymbol{\eta} : \mathcal{X} \to \Delta^K$, where the *j*-th component of the vector is $[\boldsymbol{\eta}(x)]_j = \mathbb{P}(Y = j \mid X = x)$. A classifier $f : \mathcal{X} \to \mathcal{Y}$ maps an instance x to a class $f(x) \in \mathcal{Y}$. Denote the risk of a classifier f with respect to distribution P_{XY} as 64 65 66 $R(f) = \mathbb{E}_{(X,Y)\sim P_{XY}} \left[\mathbb{1}_{\{f(X)\neq Y\}}\right]$. The Bayes optimal classifier for P_{XY} is $f^*(x) \in \arg \max \eta(x)$. The Bayes risk, which is the minimum achievable risk, is then denoted as $R^* = R(f^*) = \inf_f R(f)$. 67 68

We consider the setting where, instead of the true class label Y, a noisy label \hat{Y} is observed. 69 The training data $(X_1, \widetilde{Y}_1), \ldots, (X_n, \widetilde{Y}_n)$ can be viewed as an i.i.d. sample drawn from a "noisy" 70 distribution $P_{X\widetilde{Y}}$. We define $P_{\widetilde{Y}|X=x}$, $\widetilde{\eta}$, \widetilde{R} and \widetilde{f}^* analogously to the "clean" distribution P_{XY} . 71

We view (X, Y, \tilde{Y}) as a jointly distributed triple. Given $(X, Y) \sim P_{XY}$, the distribution of the noisy 72

label \widetilde{Y} is determined by the matrix-valued function $E : \mathcal{X} \to \{M \in \mathbb{R}^{K \times K} : M \text{ is row stochastic}\},\$ 73

which is called the *noise transition matrix*, and whose (i, j)-th element is 74

$$[\boldsymbol{E}(x)]_{i,j} = \mathbb{P}\left(\widetilde{Y} = j \mid Y = i, X = x\right).$$

Note that the noisy and clean class posteriors satisfy $\tilde{\eta}(x) = E(x)^{\top} \eta(x)$, with $^{\top}$ denoting transpose. 75

The goal of learning from label noise is to find a classifier that is able to minimize the "clean test 76 error," that is, the risk R defined w.r.t. P_{XY} , even though the learner's access is limited to corrupted 77 training data $(X_i, \tilde{Y}_i) \stackrel{\text{i.i.d.}}{\sim} P_{X\tilde{Y}}$. In essence, label noise can be conceptualized as a type of domain adaptation problem, where P_{XY} is the source domain, $P_{X\tilde{Y}}$ is the target domain, and the source 78

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and target are linked by "posterior drift", which means that the source and target have the same

81 X-marginal, but the "posteriors" (distribution of label given X) may change [Scott, 2019, Cai and 82 Wei, 2021, Maity et al., 2023]. Thus, a label noise problem is captured by a triple $(P_X, \eta, \tilde{\eta})$.

3.2 Relative signal strength

To study label noise, we introduce the concept of *relative signal strength* (RSS). This is a pointwise measure of how much "signal" (certainty about the label) is contained in the noisy distribution relative to the clean distribution. Previous work [Cannings et al., 2020, Cai and Wei, 2021] has examined a related concept within the context of binary classification, under the restriction that clean and noisy Bayes classifiers are identical. Our definition incorporates multi-class classification and relaxes the requirement that the clean and noisy Bayes classifiers agree.

90 **Definition 1 (Relative Signal Strength)** For any class probability vectors η , $\tilde{\eta}$, define the relative 91 signal strength (*RSS*) at $x \in \mathcal{X}$ as

$$\mathcal{M}(x;\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}) = \min_{j \in \mathcal{Y}} \quad \frac{\max_{i}[\boldsymbol{\eta}(x)]_{i} - [\boldsymbol{\eta}(x)]_{j}}{\max_{i}[\boldsymbol{\eta}(x)]_{i} - [\boldsymbol{\eta}(x)]_{j}},\tag{1}$$

where $0/0 := +\infty$. Furthermore, for $\kappa \in [0, \infty)$, denote the set of points whose RSS exceeds κ as

$$\mathcal{A}_{\kappa}(\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}) = \{x \in \mathcal{X} : \mathcal{M}(x; \boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}) > \kappa\}.$$

⁹³ $\mathcal{M}(x; \eta, \tilde{\eta})$ is a point-wise measure of how much "signal" the noisy posterior contains about the ⁹⁴ clean posterior. To gain some intuition, first notice that if the noisy Bayes classifier predicts a different ⁹⁵ class than the clean Bayes classifier, the RSS is 0 by taking $j = \arg \max \tilde{\eta}$ (assuming for simplicity ⁹⁶ that the arg max is a singleton set). Now suppose the clean and noisy Bayes classifiers *do* make the ⁹⁷ same prediction at *x*, say *i*^{*}, and consider a fixed *j*. If

$$\frac{[\widetilde{\boldsymbol{\eta}}(x)]_{i^*} - [\widetilde{\boldsymbol{\eta}}(x)]_j}{[\boldsymbol{\eta}(x)]_{i^*} - [\boldsymbol{\eta}(x)]_j}$$

is small, it means that the clean Bayes classifier is relatively certain that j is not the correct clean label, while the noisy Bayes classifier is less certain that j is not the correct noisy label. Taking the minimum over j gives the relative signal strength at x. As we formalize in the next section, a large RSS at x ensures that a small (pointwise) noisy excess risk at x implies a small (pointwise) clean excess risk. To gain more intuition, consider the following examples.

103 **Example 1** When $\eta(x) = \begin{bmatrix} 0 & 1 & 0 \end{bmatrix}^{\top}$ and $\tilde{\eta}(x) = \begin{bmatrix} 0.3 & 0.6 & 0.1 \end{bmatrix}^{\top}$,

$$\mathcal{M}(x;\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}) = \min_{j\in\mathcal{Y}} \quad \frac{\max_i[\widetilde{\boldsymbol{\eta}}(x)]_i - [\widetilde{\boldsymbol{\eta}}(x)]_j}{\max_i[\boldsymbol{\eta}(x)]_i - [\boldsymbol{\eta}(x)]_j} = \frac{[\widetilde{\boldsymbol{\eta}}(x)]_2 - [\widetilde{\boldsymbol{\eta}}(x)]_1}{[\boldsymbol{\eta}(x)]_2 - [\boldsymbol{\eta}(x)]_1} = \frac{0.6 - 0.3}{1 - 0} = 0.3.$$

Here, the clean Bayes classifier is absolutely certain about its prediction, while the noisy Bayes
 classifier is much less certain.

106 **Example 2** When $\boldsymbol{\eta}(x) = \begin{bmatrix} 0 & 1 & 0 \end{bmatrix}^{\top}$ and $\widetilde{\boldsymbol{\eta}}(x) = \begin{bmatrix} 0 & 0 & 1 \end{bmatrix}^{\top}$, $\mathcal{M}(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) = \min_{j \in \mathcal{Y}} \quad \frac{\max_i [\widetilde{\boldsymbol{\eta}}(x)]_i - [\widetilde{\boldsymbol{\eta}}(x)]_j}{\max_i [\boldsymbol{\eta}(x)]_i - [\boldsymbol{\eta}(x)]_j} = \frac{[\widetilde{\boldsymbol{\eta}}(x)]_3 - [\widetilde{\boldsymbol{\eta}}(x)]_3}{[\boldsymbol{\eta}(x)]_2 - [\boldsymbol{\eta}(x)]_3} = \frac{1-1}{1-0} = 0.$

¹⁰⁷ The zero signal strength results from $\tilde{\eta}$ and η leading to different predictions.

Example 3 (Comparison to KL divergence) When $\eta(x) = [0.05 \ 0.7 \ 0.25]^{\top}$, and $\tilde{\eta}^{(1)}(x) = [0.25 \ 0.7 \ 0.05]^{\top}$, $\tilde{\eta}^{(2)}(x) = [0.1 \ 0.6 \ 0.3]^{\top}$,

$$\frac{1}{\mathcal{D}_{\mathrm{KL}}\left(\boldsymbol{\eta} \, \big\| \, \widetilde{\boldsymbol{\eta}}^{(1)}\right)} < \frac{1}{\mathcal{D}_{\mathrm{KL}}\left(\boldsymbol{\eta} \, \big\| \, \widetilde{\boldsymbol{\eta}}^{(2)}\right)} \quad \text{while} \quad \mathcal{M}\left(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}^{(1)}\right) > \mathcal{M}\left(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}^{(2)}\right).$$

Here, $\tilde{\eta}^{(2)}$ is "closer" to η in terms of KL divergence, but $\tilde{\eta}^{(1)}$ provides more information in terms of

predicting the $\arg \max of \eta$. There is no conflict: KL divergence considers the similarity between

two (whole) distributions, while the task of classification only focuses on predicting the arg max.

113 This also illustrates why our notion of RSS is better suited for the label noise problem than other

114 general-purpose distance measures between distributions.

- A desirable learning scenario would be if $\mathcal{A}_{\kappa}(\eta, \tilde{\eta}) = \mathcal{X}$ for some large κ , indicating that the signal 115
- strength is big across the entire space. Unfortunately, this ideal situation is generally not achievable. 116
- To understand this limitation, we begin by making a mild assumption about the label noise: 117

$$\left| \arg \max \boldsymbol{\eta}(x) \right| \le \left| \arg \max \widetilde{\boldsymbol{\eta}}(x) \right|, \quad \text{almost surely } (P_X)$$
 (C)

where $|\cdot|$ denotes set cardinality, that is, that $\arg \max \widetilde{\eta}(x)$ contains at least as many entries as 118 $\arg \max \eta$. In other words, when there are ties for the clean Bayes prediction, the noisy Bayes 119 prediction has at least as many ties. This assumption is reasonable because label noise typically 120 introduces ambiguity by "confusing" the class probabilities, resulting in more ties.

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Proposition 1 If (C) holds, then $\mathcal{A}_0(\eta, \tilde{\eta}) = \{x \in \mathcal{X} : \arg \max \tilde{\eta}(x) = \arg \max \eta(x)\}.$ 122

This proposition, which is proved in Appendix A.1.1, indicates that A_0 , the region with positive 123 RSS, precisely corresponds to the area where the true and noisy Bayes classifiers agree. Accordingly, 124 $\mathcal{X} \setminus \mathcal{A}_0$, the zero signal region, is the region where Bayes decision rules differ. The "region of strong 125 signal," \mathcal{A}_{κ} , is a subset of \mathcal{A}_0 . Since the clean and noisy Bayes classifiers will typically disagree for 126 at least some $x, A_0 \neq \mathcal{X}$ in general. We note that the (somewhat strong) assumption that $A_0 = \mathcal{X}$ 127 has been made in prior studies [Cannings et al., 2020, Cai and Wei, 2021]. 128

Posterior Drift Model Class. Now putting definitions together, we consider the posterior drift model 129 Π defined over the triple $(P_X, \eta, \tilde{\eta})$. Let $\epsilon \in [0, 1], \kappa \in (0, +\infty)$, and define 130

$$\Pi(\epsilon,\kappa) := \Big\{ (P_X, \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) : P_X \Big(\mathcal{A}_{\kappa} (\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) \Big) \ge 1 - \epsilon \Big\}.$$

This is a set of triples (label noise problems) such that \mathcal{A}_{κ} , the region with RSS at least κ , covers at 131 least $1 - \epsilon$ of the probability mass. In the next section, we will demonstrate that the performance 132 within \mathcal{A}_{κ} can be guaranteed, whereas learning outside the region \mathcal{A}_{κ} is provably challenging. 133

Upper and lower bounds 4 134

In this section, we establish both upper and lower bounds for excess risk under multi-class instance-135 dependent label noise. 136

4.1 Minimax lower bound 137

Our first theorem reveals a fundamental limit: no classifier trained using noisy data can surpass the 138 constraints imposed by relative signal strength in a minimax sense. To state the theorem, we employ 139 the following notation and terminology. Denote the noisy training data by $Z^n = \{(X_i, \widetilde{Y}_i)\}_{i=1}^n \overset{i.i.d.}{\sim}$ 140 $P_{X\tilde{Y}}$. A *learning rule* \hat{f} is an algorithm that takes Z^n and outputs a classifier. The risk $R(\hat{f})$ of a learning rule is a random variable, where the randomness is due to the draw Z^n . 141 142

Theorem 1 (Minimax Lower Bound) Let $\epsilon \in [0, 1], \kappa > 0$. Then 143

$$\inf_{\hat{f}} \sup_{(P_X, \boldsymbol{\eta}, \boldsymbol{\tilde{\eta}}) \in \Pi(\epsilon, \kappa)} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \geq \frac{K - 1}{K} \epsilon + \Omega\left(\frac{1}{\kappa} \sqrt{\frac{1}{n}}\right),$$

where the inf is over all learning rules. 144

Proof Sketch. The idea is to pick a jointly distributed triple (X, Y, \tilde{Y}) such that: Y is independent of 145 \widetilde{Y} for $x \in \mathcal{X} \setminus \mathcal{A}_{\kappa}$ and Y and \widetilde{Y} are supported on the same classes for $x \in \mathcal{A}_{\kappa}$. Consider the two 146 regions separately: the excess risk on $\mathcal{X} \setminus \mathcal{A}_{\kappa}$ is no less than $\frac{K-1}{K}\epsilon$, while the excess risk on \mathcal{A}_{κ} decreases with sample size n, and the rate is affected by the RSS κ . See Appendix A.1.2. 147 148

The proof offers insights into how label noise impacts the learning process: if the signal is low, 149 learning is hopeless, and the learner incurs an irreducible error of $(1 - 1/K)\epsilon$; in the high signal 150 region, the signal strength κ determines the rate of convergence. These aspects determine fundamental 151 limits that no classifier trained only on noisy data can overcome without additional assumptions. 152

153 4.2 Upper bound

This subsection establishes an upper bound for NI-ERM, the empirical risk minimizer trained on noisy data. This result implies that NI-ERM is minimax optimal, a potentially surprising result given that NI-ERM is arguably the simplest approach one might consider.

Theorem 2 (Excess Risk Upper Bound of NI-ERM) Let $\epsilon \in [0,1], \kappa > 0$. Consider any ($P_X, \eta, \tilde{\eta}$) $\in \Pi(\epsilon, \kappa)$, assume function class \mathcal{F} has Natarajan dimension V, and the noisy Bayes classifier \tilde{f}^* belongs to \mathcal{F} . Let $\hat{f} \in \mathcal{F}$ be the ERM trained on $Z^n = \{(X_i, \tilde{Y}_i)\}_{i=1}^n$. Then for n > 2V,

$$\mathbb{E}_{Z^n}\left[R\left(\hat{f}\right) - R(f^*)\right] \le \epsilon + \mathcal{O}\left(\frac{1}{\kappa}\sqrt{\frac{V}{n}}\right).$$

Natarajan dimension is a multiclass analogue of VC dimension. The upper bound (Theorem 2)
matches the lower bound (which remains unchanged under the setting of Theorem 2 – see Appendix
A.1.2 for a refined statement) in terms of both the irreducible error and the rate of convergence. This
result is surprising as it indicates that the simplest possible approach, which ignores the presence of
noise, is optimal. No learning rule could perform significantly better in this minimax sense.

166 5 Practical implication

The modern practice of machine learning often involves training a deep neural network. In complex tasks involving noisy labels, the naïve NI-ERM is often outperformed by state-of-the-art methods by a significant extent [Li et al., 2020, Xiao et al., 2023]. This is consistent with the finding that directly training a large neural network on noisy data frequently leads to overfitting [Zhang et al., 2021a].

Yet this is not grounds for abandoning NI-ERM altogether as a practical strategy. Instead of using
 NI-ERM for end-to-end training of a deep neural network, we instead propose the following simple,
 two-step procedure, termed 'feature extraction + NI-ERM'.

Perform feature extraction using any method (e.g., transfer learning or self-supervised learning)
 that does not require labels.

Learn a simple classifier (e.g., a linear classifier) on top of these extracted features, using the
 noisily labelled data, in a noise-ignorant way.

We acknowledge the that practical idea of ignoring label noise is not new [Ghosh and Lan, 2021],
but the full power of this approach has not been previously recognized. For example, prior work
that has suggested ignoring the label noise usually augments this approach with additional heuristics
[Zheltonozhskii et al., 2022, Xue et al., 2022].

Remarkably, this two-step approach attains extremely strong performance. We conducted experiments 182 on the CIFAR image data under two scenarios: synthetic label flipping and realistic human label 183 errors [Wei et al., 2022], as shown in Figure 1. We examine three different feature extractors: the 184 DINOv2 foundation model [Oquab et al., 2023], ResNet-50 features extracted from training on 185 ImageNet [He et al., 2016], and self-supervised ResNet-50 using contrastive loss [Chen et al., 2020]. 186 We also compared to a simple linear model trained on the raw pixel intensities, and a ResNet-50 187 trained end-to-end. We observed that ResNet-50 exhibits a "tempered overfitting" behavior, consistent 188 with previous findings [Zhang et al., 2021a, Mallinar et al., 2022]. The linear model demonstrates 189 robustness to noise, but suffers from significant approximation error. 190

Conversely, the FE+NI-ERM approach enjoys the best of both worlds. Regardless of how the feature 191 extraction is carried out, the resulting models exhibit robustness to label noise, while the overall 192 accuracy depends entirely on the quality of the extracted features. This is illustrated in Figure 1, 193 where the flatness of the accuracy curves as noise increases indicates the robustness, and the intercept 194 at zero label noise is a measure of the feature quality. Importantly, this property holds true even under 195 realistic label noise of CIFAR-N [Wei et al., 2022]. In fact, we find that using the DINOv2 [Oquab 196 et al., 2023] extracted features in our FE+NI-ERM approach yields state of the art results on the 197 CIFAR-10N and CIFAR-100N benchmarks, across the noise levels, as shown in Table 1. 198



Figure 1: A linear model trained on features obtained from either transfer learning (pretrained ResNet-50 on ImageNet [He et al., 2016]), self-supervised learning (ResNet-50 trained on CIFAR-10 images with contrastive loss [Chen et al., 2020]), or a pretrained self-supervised foundation model DINOv2 [Oquab et al., 2023] significantly boosts the performance of the original linear model. In contrast, directly training a ResNet-50 leads to overfitting.

We reiterate that the only hyperparameters of our model are the hyperparameters of the linear classifier, which are tuned automatically using standard cross-validation on the noisy labels. This in contrast to the implementations of many methods on the CIFAR-N leaderboard (http://noisylabels.com/), where the hyperparameters are hard-coded. Altogether, this dominant performance, along with the simplicity of the approach and the lack of any untunable hyperparameters, suggest that the

the simplicity of the approach and the lack of any untunable hyperparameters, suggest th FE+NI-ERM is very powerful, and indicates a need for further investigation of its properties.

Table 1: Performance comparison with CIFAR-N leaderboard (http://noisylabels.com/) in terms of testing accuracy. "Aggre", "Rand1", ..., "Noisy" denote various types of human label noise. We compare with four methods that covers the top three performance for all noise categories: ProMix [Xiao et al., 2023], ILL [Chen et al., 2023], PLS [Albert et al., 2023] and DivideMix [Li et al., 2020]. Our approach, a Noise Ignorant linear model trained on features extracted by the self-supervised foundation model DINOv2 [Oquab et al., 2023] achieves new state-of-the-art results, highlighted in bold. We employed Python's sklearn logistic regression and cross-validation functions without data augmentation; the results are deterministic and directly reproducible.

Leaderboard			CIFAR-10N			CIFAR-100N
Methods	Aggre	Rand1	Rand2	Rand3	Worst	Noisy
ProMix	97.65 ± 0.19	97.39 ± 0.16	97.55 ± 0.12	97.52 ± 0.09	$\textbf{96.34} \pm \textbf{0.23}$	73.79 ± 0.28
ILL	96.40 ± 0.03	96.06 ± 0.07	95.98 ± 0.12	96.10 ± 0.05	93.55 ± 0.14	68.07 ± 0.33
PLS	96.09 ± 0.09	95.86 ± 0.26	95.96 ± 0.16	96.10 ± 0.07	93.78 ± 0.30	73.25 ± 0.12
DivideMix	95.01 ± 0.71	95.16 ± 0.19	95.23 ± 0.07	95.21 ± 0.14	92.56 ± 0.42	71.13 ± 0.48
FE + NI-ERM	$\textbf{98.69} \pm \textbf{0.00}$	$\textbf{98.80} \pm \textbf{0.00}$	$\textbf{98.65} \pm \textbf{0.00}$	$\textbf{98.67} \pm \textbf{0.00}$	95.71 ± 0.00	$\textbf{83.17} \pm \textbf{0.00}$

205 6 Conclusions

This work presents a rigorous theory for learning under multi-class, instance-dependent label noise. We establish matching upper and lower bounds for excess risk . Our theory reveals the minimax optimality of Noise Ignorant Empirical Risk Minimizer (NI-ERM). To make this theory practical, we provide a simple modification leveraging a feature extractor with NI-ERM, demonstrating significant performance enhancements. A limitation of this work is that our methodology warrants more extensive experimental evaluation.

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A Appendix / supplemental material 315

- A.1 Proofs 316
- A.1.1 Proof of Proposition 1 317
- **Proposition** Assume (C) holds. Then for $\kappa \ge 0$, 318

$$\mathcal{A}_0(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) = \left\{ x \in \mathcal{X} : \arg \max \widetilde{\boldsymbol{\eta}}(x) = \arg \max \boldsymbol{\eta}(x) \right\} \quad almost \ surely.$$

Proof. Notice that 319

$$\mathcal{M}(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) = 0 \quad \iff \quad \arg \max \widetilde{\boldsymbol{\eta}}(x) \not\subseteq \arg \max \boldsymbol{\eta}(x).$$

- This is because $\mathcal{M}(x; \eta, \tilde{\eta}) = 0$ when the numerator is zero and the denominator is non-zero, which 320 happens when $\arg \max \widetilde{\eta}(x) \not\subseteq \arg \max \eta(x)$. An equivalent statement of this is 321

$$\mathcal{M}(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) > 0 \quad \iff \quad \arg \max \widetilde{\boldsymbol{\eta}}(x) \subseteq \arg \max \boldsymbol{\eta}(x).$$

Under assumption (C), the set $\arg \max \tilde{\eta}(x)$ cannot be a proper subset of $\arg \max \eta(x)$, and therefore 322

$$\mathcal{M}(x; \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) > 0 \quad \iff \quad \arg \max \widetilde{\boldsymbol{\eta}}(x) = \arg \max \boldsymbol{\eta}(x)$$

Thus, 323

$$\mathcal{A}_0(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}) = \left\{ x \in \mathcal{X} : \arg \max \widetilde{\boldsymbol{\eta}}(x) = \arg \max \boldsymbol{\eta}(x) \right\},$$

almost everywhere. 324

A.1.2 Proof of Lower Bound: Theorem 1 325

Now we provide a more formal statement of the minimax lower bound and its proof. We begin with 326 the scenario where the noisy distribution $P_{X\widetilde{Y}}$ has zero Bayes risk as an introductory example. The 327 proof for the general case follows a similar strategy but involves more complex bounding techniques. 328 We recommend that interested readers first review the proof of the zero-error version to build a solid 329 understanding before tackling the general case. 330

Now consider a more restricted set of $\Pi(\epsilon, \kappa)$: 331

$$\Pi(\epsilon,\kappa,V,0) := \Big\{ (P_X,\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}) : P_X\Big(\mathcal{A}_\kappa(\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}})\Big) \ge 1-\epsilon, P_X \text{ supported on } V+1 \text{ points}, \widetilde{R}^* = 0 \Big\}.$$

Theorem (Minimax Lower Bound: when $\widetilde{R}^* = 0$) Let $\epsilon \in [0, 1], \kappa > 0, V > 1$. For any learning rule \widehat{f} based upon $Z^n = \{(X_i, \widetilde{Y}_i)\}_{i=1}^n$, and $n > \max(V - 1, 2)$, 332 333

$$\begin{split} \sup_{\substack{(P_X, \boldsymbol{\eta}, \tilde{\boldsymbol{\eta}}) \in \Pi(\epsilon, \kappa)}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] &\geq \sup_{\substack{(P_X, \boldsymbol{\eta}, \tilde{\boldsymbol{\eta}}) \in \Pi(\epsilon, \kappa, V, 0)}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \\ &\geq \frac{K - 1}{K} \epsilon + \frac{1}{\kappa} \frac{(V - 1)(1 - \epsilon)}{8en} \end{split}$$

Proof. Consider any V + 1 distinct points x_0, x_1, \ldots, x_V . Choose

$$P_X(x) = \begin{cases} \epsilon & x = x_0\\ (1-\epsilon) \cdot \frac{1}{n} & x = x_1, \dots, x_{V-1},\\ (1-\epsilon) \cdot \left(1 - \frac{V-1}{n}\right) & x = x_V. \end{cases}$$

where n > V - 1, and define the clean and noisy class posteriors by

If
$$x = x_0$$
, then $\eta(x) = e_j$, $\tilde{\eta}(x) = e_1$, $j \in \{1, 2, \dots K\}$ (2)
If $x = x_t$, $1 \le t \le V - 1$, then $\eta(x) = \begin{bmatrix} \frac{1}{2} + \frac{1}{2(\kappa+\delta)} \cdot (-1)^{b_t+1} \\ \frac{1}{2} - \frac{1}{2(\kappa+\delta)} \cdot (-1)^{b_t+1} \\ 0 \\ \vdots \\ 0 \end{bmatrix}$, $\tilde{\eta}(x) = e_{b_t}, b_t \in \{1, 2\}, \delta > 0$
(3)

If
$$x = x_V$$
, then $\boldsymbol{\eta}(x) = \begin{bmatrix} \frac{1}{2} + \frac{1}{2(\kappa+\delta)} \\ \frac{1}{2} - \frac{1}{2(\kappa+\delta)} \\ 0 \\ \vdots \\ 0 \end{bmatrix}$, $\widetilde{\boldsymbol{\eta}}(x) = \boldsymbol{e}_1$, (4)

336 where e_i denotes the one-hot vector whose *i*-th element is one.

337 The triple $(P_X, \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}})$ is thus parameterized by $j, \boldsymbol{b} := [b_1 \ b_2 \ \cdots \ b_{V-1}]^\top$, and δ .

This construction ensures $(P_X, \eta, \tilde{\eta}) \in \Pi(\epsilon, \kappa, V, 0)$. In particular,

$$\begin{array}{ll} \mathcal{A}_{\kappa} \supseteq \{x_1, x_2, \dots, x_V\}, & P_X(\mathcal{A}_{\kappa}) \ge 1 - \epsilon, \\ \mathcal{X} \setminus \mathcal{A}_{\kappa} \subseteq \{x_0\}, & P_X(\mathcal{X} \setminus \mathcal{A}_{\kappa}) \le \epsilon, \end{array}$$

and $\widetilde{R}^* = 0$ because $\widetilde{\eta}(x)$ is one-hot for all x.

For classifier any f, by definition, its risk equals to

$$R(f) = \mathbb{E}_{X,Y} \left[\mathbb{1}_{f(X)\neq Y} \right]$$

= $\mathbb{E}_X \mathbb{E}_{Y|X} [\mathbb{1}_{f(X)\neq Y}]$
= $\mathbb{E}_X \mathbb{E}_{Y|X} [1 - \mathbb{1}_{f(X)=Y}]$
= $\mathbb{E}_X \left[1 - [\boldsymbol{\eta}(X)]_{f(X)} \right]$
= $\int_{\mathcal{X}} \left(1 - [\boldsymbol{\eta}(x)]_{f(x)} \right) dP_X(x),$

Under our construction of P_X , R(f) can be decomposed into two parts

$$R(f) = \underbrace{\int_{\{x_0\}} \left(1 - [\eta(x)]_{f(x)}\right) dP_X(x)}_{:=R_0(f)} + \underbrace{\int_{\{x_1,\dots,x_V\}} \left(1 - [\eta(x)]_{f(x)}\right) dP_X(x)}_{:=R_V(f)},$$

342 so does the excess risk

$$R(f) - R(f^*) = \left(R_0(f) - R_0(f^*)\right) + \left(R_V(f) - R_V(f^*)\right).$$

Recall that in our construction, $(P_X, \eta, \tilde{\eta})$ is parameterized by j, b, and δ . Therefore

$$\sup_{\substack{(P_X,\boldsymbol{\eta},\tilde{\boldsymbol{\eta}})\in\Pi(\epsilon,\kappa,V,0)}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \ge \sup_{j,\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right]$$
$$= \sup_{j,\boldsymbol{b},\delta} \left\{ \mathbb{E}_{Z^n} \left[R_0\left(\hat{f}\right) - R_0(f^*) \right] \right\}$$
$$+ \mathbb{E}_{Z^n} \left[R_V\left(\hat{f}\right) - R_V(f^*) \right] \right\}$$
$$= \sup_{j} \mathbb{E}_{Z^n} \left[R_0\left(\hat{f}\right) - R_0(f^*) \right]$$
$$+ \sup_{\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[R_V\left(\hat{f}\right) - R_V(f^*) \right]$$

- where the last equality holds because region $\{x_0\}$ only depends on j, while region $\{x_1, \ldots, x_V\}$
- only depends on \boldsymbol{b}, δ .
- ³⁴⁶ In the remaining part of the proof, we will examine

$$\sup_{j} \mathbb{E}_{Z^{n}} \left[R_{0}\left(\hat{f}\right) - R_{0}(f^{*}) \right]$$
(5)

347 and

$$\sup_{\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[R_V \left(\hat{f} \right) - R_V(f^*) \right]$$
(6)

348 separately.

Let's start with the first term (5), which acts over the "low signal strength" region $\{x_0\}$. Since η is one-hot on $\{x_0\}$, its Bayes risk over that is zero

$$\sup_{j} \mathbb{E}_{Z^{n}} \left[R_{0}\left(\hat{f}\right) - R_{0}(f^{*}) \right] = \sup_{j} \mathbb{E}_{Z^{n}} \left[R_{0}\left(\hat{f}\right) \right]$$
$$= \sup_{j} \mathbb{E}_{Z^{n}} \left[\int_{\{x_{0}\}} \mathbb{1}_{\hat{f}(x)\neq j} dP_{X}(x) \right].$$

To deal with \sup_j , we use a technique called "the probabilistic method": replace j with a random variable $J \sim \text{Uniform}\{1, 2, \dots, K\}$:

$$\sup_{j} \mathbb{E}_{Z^{n}} \left[\int_{\{x_{0}\}} \mathbb{1}_{\hat{f}(x)\neq j} dP_{X} \right] \geq \mathbb{E}_{J, Z^{n}} \left[\int_{\{x_{0}\}} \mathbb{1}_{\hat{f}(x)\neq J} dP_{X}(x) \right]$$
$$= \mathbb{E}_{Z^{n}} \left[\mathbb{E}_{J|Z^{n}} \left[\int_{\{x_{0}\}} \mathbb{1}_{\hat{f}(x)\neq J} dP_{X}(x) \right] \right].$$

- Again, notice that J is an independent draw. Even if the point x_0 is observed in Z^n , the associated
- noisy label $\tilde{Y} = 1$ does not give any information about the clean label Y = J. Thus

$$\mathbb{E}_{Z^n} \left[\mathbb{E}_{J|Z^n} \left[\int_{\{x_0\}} \mathbb{1}_{\hat{f}(x) \neq J} dP_X(x) \right] \right] = \mathbb{E}_{Z^n} \left[\mathbb{E}_J \left[\int_{\{x_0\}} \mathbb{1}_{\hat{f}(x) \neq J} dP_X(x) \right] \right]$$
$$= \mathbb{E}_{Z^n} \left[\int_{\{x_0\}} \mathbb{E}_J \left[\mathbb{1}_{\hat{f}(x) \neq J} \right] dP_X(x) \right]$$
$$= \mathbb{E}_{Z^n} \left[\int_{\{x_0\}} \left(1 - \frac{1}{K} \right) dP_X(x) \right]$$
$$= \left(1 - \frac{1}{K} \right) \epsilon.$$

Now we have the minimax lower bound for the first part (5):

$$\sup_{j} \mathbb{E}_{Z^{n}} \left[R_{\{x_{0}\}} \left(\hat{f} \right) - R_{\{x_{0}\}} (f^{*}) \right] \geq \left(1 - \frac{1}{K} \right) \epsilon.$$

For the second part (6), which is over $\{x_1, \ldots, x_V\}$, due to the relative signal strength condition, and

from our explicit construction in Eqn. (3) and (4), the excess risk w.r.t. true and noisy distribution are related by

$$\begin{aligned} R_V(f) - R_V(f^*) &= \int_{\{x_1, \dots, x_V\}} \left(\max \boldsymbol{\eta}(x) - [\boldsymbol{\eta}(x)]_{f(x)} \right) dP_X(x) \\ &= \int_{\{x_1, \dots, x_V\}} \frac{1}{\kappa + \delta} \Big(\max \widetilde{\boldsymbol{\eta}}(x) - [\widetilde{\boldsymbol{\eta}}(x)]_{f(x)} \Big) dP_X(x) \quad \because \text{ by construction of } \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \\ &= \frac{1}{\kappa + \delta} \left(\widetilde{R}_V(f) - \widetilde{R}_V(\widetilde{f}^*) \right), \end{aligned}$$

where $\widetilde{R}_V(f) := \int_{\{x_1,\dots,x_V\}} \left(1 - [\widetilde{\boldsymbol{\eta}}(x)]_{f(x)}\right) dP_X(x)$. Also note that $f^*(x) = \widetilde{f}^*(x)$ for $x \in \{x_1,\dots,x_V\}$, which is a result of our construction of $\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}}$.

361 Then

$$\sup_{\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[R_V\left(\hat{f}\right) - R_V(f^*) \right] = \sup_{\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_V(f) - \widetilde{R}_V(\widetilde{f}^*) \right) \right].$$

This allows us to reduce the label noise problem into a standard learning problem: we have an iid sample Z^n from $P_{X\tilde{Y}}$ and consider the risk evaluated on the same distribution $P_{X\tilde{Y}}$. The remainder of the proof is similar to the proof of Theorem 14.1 in Devroye et al. [1996].

Notice that by our construction, \tilde{Y} is a deterministic function of X. To be specific, $\tilde{Y} = \tilde{f}^*(X)$, where

$$\tilde{f}^*(x) = \begin{cases} 1 & x = x_0, \\ b_t & x = x_t, \ 1 \le t \le V - 1 \\ 1 & x = x_V \end{cases}$$

- 367 is the noisy Bayes classifier.
- We use the shorthand $f_{b} := \tilde{f}^{*}$ to denote that the noisy Bayes classifier depends on b, we also write the learning rule $\hat{f}(x; Z^{n}) := \hat{f}(x)$ to indicate the dependence of the learning rule \hat{f} on the random
- 370 sample Z^n .
- 371 Since the noisy Bayes risk is zero,

$$\sup_{\boldsymbol{b},\boldsymbol{\delta}} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_V(f) - \widetilde{R}_V(\widetilde{f}^*) \right) \right] = \sup_{\boldsymbol{b},\boldsymbol{\delta}} \frac{1}{\kappa + \delta} \mathbb{E}_{Z^n} \left[\widetilde{R}_V(\widehat{f}) \right].$$

Again, use the probabilistic method, replace **b** with $\mathbf{B} \sim \text{Uniform}\{1,2\}^{V-1}$,

$$\sup_{\boldsymbol{b},\delta} \frac{1}{\kappa+\delta} \mathbb{E}_{Z^n} \left[\widetilde{R}_V(\hat{f}) \right] \ge \sup_{\delta} \frac{1}{\kappa+\delta} \mathbb{E}_{\boldsymbol{B},Z^n} \left[\widetilde{R}_V(\hat{f}) \right]$$
$$= \sup_{\delta} \frac{1}{\kappa+\delta} \mathbb{E}_{Z^n} \left[\mathbb{E}_{\boldsymbol{B}|Z^n} \left[\int_{\{x_1,\dots,x_V\}} \mathbb{1}_{\hat{f}(x)\neq f_{\boldsymbol{B}}(x)} dP_X(x) \right] \right]$$

Since we have $B \sim \text{Uniform}\{1,2\}^{V-1}$ and also $Z|B \sim P_{X\tilde{Y}}$, then by Bayes rule (or eye-balling), we get the posterior distribution of $B|Z^n$, to be specific:

$$\forall x \in \{x_1, \cdots, x_V\}, \ \mathbb{P}\left(f_{\boldsymbol{B}}(x) = 1 | Z^n\right) = \begin{cases} \frac{1}{2} & x \neq X_1, \ldots, x \neq X_n, x \neq x_V\\ 0 \text{ or } 1 & \text{otherwise,} \end{cases}$$

- where we overload the notation \mathbb{P} to denote conditional probability of $B|Z^n$.
- Then the optimal decision rule for predicting B based on sample Z^n is:

$$f^*_{\boldsymbol{B}}(x; Z^n) := \begin{cases} \widetilde{Y}_i & x = X_i, i \in \{1, 2, \dots, n\} \\ 1 & x = x_V \\ \text{random guess from } \{1, 2\} & x \neq X_1, \dots, x \neq X_n, x \neq x_V. \end{cases}$$

Therefore, roughly speaking, the error comes from the probability of $X \in \{x_1, \ldots, x_V\}$ not being one of observed X_i : for any \hat{f} ,

$$\begin{split} \mathbb{E}_{\boldsymbol{B},Z^n}\left[\tilde{R}_V(\hat{f})\right] &= \mathbb{E}_{Z^n}\left[\mathbb{E}_{\boldsymbol{B}|Z^n}\left[\int_{\{x_1,\dots,x_V\}} \mathbbm{1}_{\hat{f}(x)\neq f_{\boldsymbol{B}}(x)} dP_X(x)\right]\right] \\ &\geq \mathbb{P}\left(X \in \{x_1,\dots,x_V\}, f_{\boldsymbol{B}}^*(X;Z^n) \neq f_{\boldsymbol{B}}(X)\right) \quad \because \text{ error of } \hat{f} > \text{ error of } f_{\boldsymbol{B}}^* \\ &= \left(1 - \frac{1}{2}\right)\mathbb{P}\left(X \neq X_1,\dots,X \neq X_n, X \neq x_V, X \in \{x_1,\dots,x_V\}\right) \\ &= \frac{1}{2}\sum_{t=1}^V \mathbb{P}\left(X \neq X_1,\dots,X \neq X_n, X \neq x_V, X = x_t\right) \\ &= \frac{1}{2}\sum_{t=1}^V \mathbb{P}\left(X_1 \neq x_t,\dots,X_n \neq x_t, x_V \neq x_t, X = x_t\right) \quad \because \text{ replace all } X \text{ with } x_t \\ &= \frac{1}{2}\sum_{t=1}^{V-1} \mathbb{P}\left(X_1 \neq x_t,\dots,X_n \neq x_t, X = x_t\right) \\ &= \frac{1}{2}\sum_{t=1}^{V-1} \mathbb{P}\left(X_1 \neq x_t,\dots,X_n \neq x_t|X = x_t\right) \mathbb{P}\left(X = x_t\right) \\ &= \frac{1}{2}\sum_{t=1}^{V-1} (1 - \mathbb{P}\left(X = x_t\right))^n \mathbb{P}\left(X = x_t\right) \\ &= \frac{1}{2}\left(V - 1\right)\left(1 - \frac{1 - \epsilon}{n}\right)^n \left(\frac{1 - \epsilon}{n}\right) \\ &= \frac{(V - 1)(1 - \epsilon)}{2n} \left(1 - \frac{1 - \epsilon}{n}\right)^{1 + \epsilon} \left(1 - \frac{1 - \epsilon}{n}\right)^{n - 1 - \epsilon} \\ &\geq \frac{(V - 1)(1 - \epsilon)}{2n} \left(1 - \frac{1}{n}\right)^2 e^{-1} \quad \because \epsilon \in [0, 1] \\ &\geq \frac{(V - 1)(1 - \epsilon)}{2n} \frac{e^{-1}}{4} = \frac{(V - 1)(1 - \epsilon)}{8en}. \quad \text{ take } n > 2 \end{split}$$

Now we get the minimax risk for the second part (6)

$$\sup_{\boldsymbol{b},\delta} \mathbb{E}_{Z^n} \left[R_{\mathcal{A}_{\kappa}} \left(\hat{f} \right) - R_{\mathcal{A}_{\kappa}}(f^*) \right] \ge \sup_{\delta} \frac{1}{\kappa + \delta} \frac{(V-1)(1-\epsilon)}{8en} \\ \ge \frac{1}{\kappa} \frac{(V-1)(1-\epsilon)}{8en} \quad \text{let } \delta \downarrow 0$$

380 Combine the two parts together, we get the final result, for $n > \max(V-1,2)$

$$\sup_{(P_X,\boldsymbol{\eta},\tilde{\boldsymbol{\eta}})\in\Pi(\epsilon,\kappa,V,0)} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \ge \frac{K-1}{K} \epsilon + \frac{1}{\kappa} \frac{(V-1)(1-\epsilon)}{8en}.$$

14

381

382 As for the general version of the lower bound, now consider the set of triples:

$$\begin{split} \Pi(\epsilon,\kappa,V,L) &:= \Big\{ \left(P_X, \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) : P_X \Big(\mathcal{A}_\kappa \left(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) \Big) \geq 1 - \epsilon, \\ P_X \text{ supported on } V + 1 \text{ points}, \frac{\widetilde{R}_{\mathcal{A}_\kappa} \left(\widetilde{f}^* \right)}{P_X \Big(\mathcal{A}_\kappa \left(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) \Big)} = L \Big\}, \end{split}$$

where $\widetilde{R}_C(f) = \int_C \left(1 - [\widetilde{\boldsymbol{\eta}}(x)]_{f(x)}\right) dP_X(x).$

Theorem (Minimax Lower Bound (General Version)) Let $\epsilon \in [0, 1], \kappa > 0, V > 1, L \in (0, 1/2)$. For any learning rule \hat{f} based upon $Z^n = \{(X_i, \tilde{Y}_i)\}_{i=1}^n$, for $n \ge \frac{V-1}{2L} \max\left\{9, \frac{1}{(1-2L)^2}\right\}$

$$\begin{split} \sup_{\substack{(P_X, \boldsymbol{\eta}, \tilde{\boldsymbol{\eta}}) \in \Pi(\epsilon, \kappa)}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] &\geq \sup_{\substack{(P_X, \boldsymbol{\eta}, \tilde{\boldsymbol{\eta}}) \in \Pi(\epsilon, \kappa, V, L)}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \\ &\geq \frac{K - 1}{K} \epsilon + \frac{1 - \epsilon}{\kappa} \sqrt{\frac{(V - 1)L}{72n}} e^{-8}. \end{split}$$

386 Proof.

Consider any V + 1 distinct points x_0, x_1, \ldots, x_V . Choose

$$P_X(x) = \begin{cases} \epsilon & x = x_0\\ (1-\epsilon) \cdot p & x = x_1, \dots, x_{V-1}\\ (1-\epsilon) \cdot (1-(V-1)p) & x = x_V. \end{cases}$$

- This imposes the constraint $(V-1)p \leq 1$, which will be satisfied in the end. Notice the differ-
- ence compared to the previous zero-error proof: we place probability mass p, rather than 1/n, on x_1, \ldots, x_{V-1} .
- 391 As for the clean and noisy class probabilities, choose

If
$$x = x_0$$
, then $\eta(x) = e_j$, $\tilde{\eta}(x) = e_1$, $j \in \{1, 2, \dots k\}$ (7)
If $x = x_t, 1 \le t \le V - 1$, then $\eta(x) = \begin{bmatrix} \frac{1}{2} + \frac{c}{\kappa + \delta} \cdot (-1)^{b_t + 1} \\ \frac{1}{2} - \frac{c}{\kappa + \delta} \cdot (-1)^{b_t + 1} \\ 0 \\ \vdots \\ 0 \end{bmatrix}$, $\tilde{\eta}(x) = \begin{bmatrix} \frac{1}{2} + c \cdot (-1)^{b_t + 1} \\ \frac{1}{2} - c \cdot (-1)^{b_t + 1} \\ 0 \\ \vdots \\ 0 \end{bmatrix}$, $b_t \in \{1, 2\}, \ \delta > 0, \ c \in \left(0, \frac{1}{2}\right)$ (8)

$$x = x_V, \quad \boldsymbol{\eta}(x) = \begin{bmatrix} \frac{1}{2} + \frac{1}{2(\kappa+\delta)} \\ \frac{1}{2} - \frac{1}{2(\kappa+\delta)} \\ 0 \\ \vdots \\ 0 \end{bmatrix}, \quad \boldsymbol{\tilde{\eta}}(x) = \boldsymbol{e}_1, \tag{9}$$

- where e_i denotes the one-hot vector whose *i*-th element is one.
- ³⁹³ The construction for class posterior is also similar to the previous proof, except that for x =

394 $x_t, t \in \{1, \dots, V-1\}, \widetilde{\boldsymbol{\eta}}$ is no longer a one-hot vector, rather has class probability separated by 2c: 395 $\left| [\widetilde{\boldsymbol{\eta}}(x)]_1 - [\widetilde{\boldsymbol{\eta}}(x)]_2 \right| = 2c$.

- The triple $(P_X, \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}})$ can be parameterized by $j, \boldsymbol{b} := [b_1 \ b_2 \ \cdots \ b_{V-1}]^\top, \delta, c \text{ and } p.$
- Again, this construction ensures $(P_X, \eta, \tilde{\eta}) \in \Pi(\epsilon, \kappa)$, to be specific:

$$\mathcal{A}_{\kappa} \supseteq \{x_1, x_2, \dots, x_V\}, \qquad P_X(\mathcal{A}_{\kappa}) \ge 1 - \epsilon, \\ \mathcal{X} \setminus \mathcal{A}_{\kappa} \subseteq \{x_0\}, \qquad P_X(\mathcal{X} \setminus \mathcal{A}_{\kappa}) \le \epsilon,$$

For any classifier f, its risk can be decomposed into two parts

$$R(f) = \underbrace{\int_{\{x_0\}} \left(1 - [\eta(x)]_{f(x)}\right) dP_X(x)}_{:=R_0(f)} + \underbrace{\int_{\{x_1,\dots,x_V\}} \left(1 - [\eta(x)]_{f(x)}\right) dP_X(x)}_{:=R_V(f)},$$

399 so does the excess risk

$$R(f) - R(f^*) = \left(R_0(f) - R_0(f^*)\right) + \left(R_V(f) - R_V(f^*)\right).$$

In our construction, $(P_X, \boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}})$ is parameterized by $j, \boldsymbol{b} := [b_1 \ b_2 \ \cdots \ b_{V-1}]^\top, \delta, c \text{ and } p$, therefore

$$\sup_{(P_X,\boldsymbol{\eta},\tilde{\boldsymbol{\eta}})\in\Pi(\epsilon,\kappa,V,L)} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \ge \sup_{j} \mathbb{E}_{Z^n} \left[R_0\left(\hat{f}\right) - R_0(f^*) \right]$$
(10)

$$+\sup_{\boldsymbol{b},\delta,c,p} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa+\delta} \left(\widetilde{R}_V(f) - \widetilde{R}_V(\widetilde{f}^*) \right) \right].$$
(11)

401 Note that we have used the fact that

$$R_V(f) - R_V(f^*) = \frac{1}{\kappa + \delta} \left(\widetilde{R}_V(f) - \widetilde{R}_V(\widetilde{f}^*) \right),$$

402 where $\widetilde{R}_{V}(f) := \int_{\{x_1,...,x_V\}} \left(1 - [\widetilde{\eta}(x)]_{f(x)}\right) dP_X(x).$

⁴⁰³ The first part (10) is exactly the same as the that of the zero-error proof, we have

$$\sup_{j} \mathbb{E}_{Z^{n}} \left[R_{0}\left(\hat{f}\right) - R_{0}(f^{*}) \right] \geq \left(1 - \frac{1}{K}\right) \epsilon.$$

- ⁴⁰⁴ From this point forward, the procedure is similar to the proof of Theorem 14.5 in Devroye et al.
- [1996]. For the second part (11), the noisy Bayes classifier is still

$$\tilde{f}^{*}(x) = \begin{cases} j & x = x_{0}, \\ b_{t} & x = x_{t}, \ 1 \le t \le V \\ 1 & x = x_{V}. \end{cases}$$

- We also use the shorthand $f_b := \tilde{f}^*$ to denote that the noisy Bayes classifier depends on b. The noisy Bayes risk is no longer zero. In fact
 - $\widetilde{R}_V(\widetilde{f}^*) = (V-1)(1-\epsilon)p\left(\frac{1}{2}-c\right)$
- ⁴⁰⁸ Notice that our construction $(P_X, \eta, \tilde{\eta}) \in \Pi(\epsilon, \kappa, V, L)$, by definition

$$\frac{\widetilde{R}_{\mathcal{A}_{\kappa}}\left(\widetilde{f}^{*}\right)}{P_{X}\left(\mathcal{A}_{\kappa}\left(\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}\right)\right)}=L$$

409 therefore

$$L = \frac{\widetilde{R}_{\mathcal{A}_{\kappa}}\left(\widetilde{f}^{*}\right)}{P_{X}\left(\mathcal{A}_{\kappa}\left(\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}}\right)\right)} \leq \frac{\widetilde{R}_{V}(\widetilde{f}^{*})}{P_{X}\left(\left\{x_{1},\ldots,x_{V}\right\}\right)} = (V-1)p\left(\frac{1}{2}-c\right),\tag{12}$$

410 where the inequality holds from $\widetilde{R}_{\mathcal{A}_{\kappa}}(\widetilde{f}^*) = \widetilde{R}_V(\widetilde{f}^*)$ and $P_X\left(\mathcal{A}_{\kappa}(\eta, \widetilde{\eta})\right) \ge P_X\left(\{x_1, \dots, x_V\}\right)$.

It should be noted that since $(V-1)p \le 1$ is required, and that c > 0, we have $L < 1 \cdot 1/2$. This is the origin of our condition L < 1/2 in the statement of the theorem. Naturally, the statement can be adjusted to min(L, 1/2) instead. 414 For fixed \boldsymbol{b} , the excess risk in region $\{x_1,\ldots,x_V\}$ becomes

$$\widetilde{R}_{V}(\widehat{f}) - \widetilde{R}_{V}(\widetilde{f}^{*}) = \int_{\{x_{1},\dots,x_{V}\}} 2c\mathbb{1}_{\widehat{f}(x)\neq f_{b}(x)} dP_{X}(x)$$
$$\geq 2c \sum_{t=1}^{V-1} (1-\epsilon)p\mathbb{1}_{\widehat{f}(x_{t})\neq f_{b}(x_{t})},$$

- where the inequality follows from the fact that we ignore the risk on point x_V .
- 416 Using the probabilistic method, replace \boldsymbol{b} with $\boldsymbol{B} \sim \text{Uniform}\{1,2\}^{V-1}$,

$$\sup_{\boldsymbol{b},\delta,c,p} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_V(\hat{f}) - \widetilde{R}_V(\tilde{f}^*) \right) \right] \ge \sup_{\delta,c,p} \mathbb{E}_{\boldsymbol{B},Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_V(\hat{f}) - \widetilde{R}_V(\tilde{f}^*) \right) \right] \\= \sup_{\delta,c,p} \frac{1}{\kappa + \delta} \mathbb{E}_{Z^n} \left[\mathbb{E}_{\boldsymbol{B}|Z^n} \left[\left(\widetilde{R}_V(\hat{f}) - \widetilde{R}_V(\tilde{f}^*) \right) \right] \right]$$

- Now, we need to calculate $B|Z^n$, which can be calculated using Bayes rule because we have B ~ Uniform $\{1,2\}^{V-1}$ and also $Z|B \sim P_{X\widetilde{Y}}$.
- To be specific, for any $x \in \{x_0, x_1, \dots, x_{V-1}\}$, assume point x_t is observed k times in training sample Z^n ,

$$\mathbb{P}(f_{B}(x) = 1 | Z^{n}) = \begin{cases} \frac{1}{2} & x \neq X_{1}, \dots, x \neq X_{n}, x \neq x_{V} \\ \mathbb{P}(B_{t} = 1 | Y_{t_{1}}, \dots, Y_{t_{k}}) & x = x_{t} = X_{t_{1}} = \dots = X_{t_{k}}, \ 1 \le t \le V - 1, \end{cases}$$

- 421 where B_t denotes the *t*-th element of vector B.
- Next we compute $\mathbb{P}(B_t = 1 | Y_{t_1} = y_1, \dots, Y_{t_k} = y_k)$ for $y_1, \dots, y_k \in \{1, 2\}$. Denote the numbers of ones and twos by $k_1 = |\{j \le k : y_j = 1\}|$ and $k_2 = |\{j \le k : y_j = 2\}|$. Using Bayes rule, we get

$$\mathbb{P}\left(B_{t}=1|Y_{t_{1}},\ldots,Y_{t_{k}}\right) = \frac{\mathbb{P}\left(B_{t}=1\cap Y_{t_{1}},\ldots,Y_{t_{k}}\right)}{\mathbb{P}\left(Y_{t_{1}},\ldots,Y_{t_{k}}\right)}$$
$$= \frac{\mathbb{P}\left(Y_{t_{1}},\ldots,Y_{t_{k}}|B_{t}=1\right)\mathbb{P}\left(B_{t}=1\right)}{\sum_{i=1}^{2}\mathbb{P}\left(Y_{t_{1}},\ldots,Y_{t_{k}}|B_{t}=i\right)\mathbb{P}\left(B_{t}=i\right)}$$
$$= \frac{(1/2+c)^{k_{1}}(1/2-c)^{k_{2}}(1/2)}{(1/2+c)^{k_{1}}(1/2-c)^{k_{2}}(1/2)+(1/2+c)^{k_{2}}(1/2-c)^{k_{1}}(1/2)}$$

⁴²⁵ After some calculation, following Theorem 14.5 in Devroye et al. [1996], we get

$$\begin{split} \sup_{\boldsymbol{b},\delta,c,p} \ \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_{\mathcal{A}_{\kappa}}(f) - \widetilde{R}_{\mathcal{A}_{\kappa}}(\tilde{f}^*) \right) \right] &\geq \sup_{\delta,c,p} \ \frac{1}{\kappa + \delta} c(V-1)(1-\epsilon) p e^{-\frac{8n(1-\epsilon)pc^2}{1-2c} - \frac{4c\sqrt{n(1-\epsilon)p}}{1-2c}} \\ &\geq \frac{1-\epsilon}{\kappa} \sup_{c,p} c(V-1) p e^{-\frac{8npc^2}{1-2c} - \frac{4c\sqrt{np}}{1-2c}} \quad \because \epsilon \geq 0, \text{ take } \delta \downarrow 0 \\ &= \frac{1-\epsilon}{\kappa} \sup_{c,p} c \ \frac{L}{1/2-c} e^{-\frac{8npc^2}{1-2c} - \frac{4c\sqrt{np}}{1-2c}} \end{split}$$

Take $c = \sqrt{\frac{(V-1)}{8nL}}$, and p is automatically specified by Eqn. (12). Then

$$\sup_{\boldsymbol{b},\boldsymbol{\delta},\boldsymbol{c},\boldsymbol{p}} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_{\mathcal{A}_{\kappa}}(f) - \widetilde{R}_{\mathcal{A}_{\kappa}}(\tilde{f}^*) \right) \right] \ge \frac{1 - \epsilon}{\kappa} \sqrt{\frac{V - 1}{8nL}} \frac{L}{1/2 - c} e^{-\frac{p(V-1) + \sqrt{\frac{4}{1 - 2c}}}{1 - 2c}}$$

Furthermore, take $n \ge \frac{9(V-1)}{2L}$, such that $c \le 1/6$. Also use the fact that $(V-1)p \le 1$, simplify the expression

$$\sup_{\boldsymbol{b},\boldsymbol{\delta},\boldsymbol{c},\boldsymbol{p}} \mathbb{E}_{Z^n} \left[\frac{1}{\kappa + \delta} \left(\widetilde{R}_{\mathcal{A}_{\kappa}}(f) - \widetilde{R}_{\mathcal{A}_{\kappa}}(\tilde{f}^*) \right) \right] \ge \frac{1 - \epsilon}{\kappa} \sqrt{\frac{V - 1}{8nL}} \frac{L}{1/2 - 1/6} e^{-\frac{1 + \sqrt{6}}{1 - 1/3}} \\ \ge \frac{1 - \epsilon}{\kappa} \sqrt{\frac{3(V - 1)L}{8n}} e^{-6}$$

In order to satisfy the condition $L \leq (V-1)p(1/2-c)$ and $(V-1)p \leq 1$, plug in $c = \sqrt{\frac{(V-1)}{8nL}}$,

- $\text{430} \quad \text{we have } n \geq \tfrac{V-1}{2L(1-2L)^2}.$
- 431 Compare two parts together

$$\sup_{\substack{(P_X, \boldsymbol{\eta}, \tilde{\boldsymbol{\eta}}) \in \Pi(\epsilon, \kappa) \\ 432}} \mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \ge \frac{K-1}{K} \epsilon + \frac{1-\epsilon}{\kappa} \sqrt{\frac{3(V-1)L}{8n}} e^{-6},$$

$$432 \quad \text{for } n \ge \frac{V-1}{2L} \max\left\{9, \frac{1}{(1-2L)^2}\right\}.$$

$$433$$

434 A.1.3 Proof of Upper Bounds: Lemma 1 and Theorem 2

Lemma 1 (Oracle Inequality under Feature-dependent Label Noise) For any $(P_X, \eta, \tilde{\eta})$ and any classifier f, we have

$$R(f) - R(f^*) \leq \inf_{\kappa} \left\{ P_X \left(\mathcal{X} \setminus \mathcal{A}_{\kappa} \left(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) \right) + \frac{1}{\kappa} \left(\widetilde{R}(f) - \widetilde{R} \left(\widetilde{f}^* \right) \right) \right\}.$$

437

- 438 *Proof.* For any $\kappa \geq 0$, the input space \mathcal{X} can be divided into two regions: $\mathcal{X} \setminus \mathcal{A}_{\kappa}$ and \mathcal{A}_{κ} .
- 439 For any f, its risk can be decomposed into two parts

$$R(f) = \mathbb{E}_{X,Y} \left[\mathbbm{1}_{f(X) \neq Y} \right]$$

$$= \mathbb{E}_X \mathbb{E}_{Y|X} [\mathbbm{1}_{f(X) \neq Y}]$$

$$= \mathbb{E}_X \mathbb{E}_{Y|X} [1 - \mathbbm{1}_{f(X) = Y}]$$

$$= \mathbb{E}_X \left[1 - [\mathbf{\eta}(X)]_{f(X)} \right]$$

$$= \int_{\mathcal{X}} \left(1 - [\mathbf{\eta}(x)]_{f(x)} \right) dP_X(x)$$

$$= \int_{\mathcal{X} \setminus \mathcal{A}_{\kappa}} \left(1 - [\mathbf{\eta}(x)]_{f(x)} \right) dP_X(x) + \int_{\mathcal{A}_{\kappa}} \left(1 - [\mathbf{\eta}(x)]_{f(x)} \right) dP_X(x)$$

440 Thereofore, excess risk equals to

$$R(f) - R(f^*) \underbrace{\int_{\mathcal{X} \setminus \mathcal{A}_{\kappa}} \left(\max \eta(x) - [\eta(x)]_{f(x)} \right) dP_X(x)}_{(a)} + \underbrace{\int_{\mathcal{A}_{\kappa}} \left(\max \eta(x) - [\eta(x)]_{f(x)} \right) dP_X(x)}_{(b)}$$

441 Now examine the two terms separately,

(a)
$$\leq \int_{\mathcal{X}\setminus\mathcal{A}_{\kappa}} 1 \, dP_X(x) = P_X\Big(\mathcal{X}\setminus\mathcal{A}_{\kappa}(\boldsymbol{\eta},\widetilde{\boldsymbol{\eta}})\Big),$$

442 and

$$(b) < \int_{\mathcal{A}_{\kappa}} \frac{1}{\kappa} \Big(\max \widetilde{\boldsymbol{\eta}}(x) - [\widetilde{\boldsymbol{\eta}}(x)]_{f(x)} \Big) dP_X(x) \quad \because \text{ by definition of relative signal strength} \leq \int_{\mathcal{X}} \frac{1}{\kappa} \Big(\max \widetilde{\boldsymbol{\eta}}(x) - [\widetilde{\boldsymbol{\eta}}(x)]_{f(x)} \Big) dP_X(x) = \frac{1}{\kappa} \Big(\widetilde{R}(f) - \widetilde{R}(\widetilde{f}^*) \Big) \quad \because \text{ by definition of } \widetilde{R}.$$

443 Since this works for any $\kappa > 0$, we then have

$$R(f) - R(f^*) \leq \inf_{\kappa > 0} \left\{ P_X \left(\mathcal{X} \setminus \mathcal{A}_{\kappa} \left(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) \right) + \frac{1}{\kappa} \left(\widetilde{R}(f) - \widetilde{R} \left(\widetilde{f}^* \right) \right) \right\}$$

444

Now, to prepare for the rate of convergence proof, we first introduce the concept of shattering in multiclass and the Natarajan dimension [Natarajan, 1989], which can be viewed as a multi-class analogy of VC dimension [Vapnik and Chervonenkis, 1971].

Definition 2 (Shattering (Multiclass)) Let \mathcal{H} be a class of functions from \mathcal{X} to $\mathcal{Y} = \{1, 2, ..., K\}$. For any set containing n distinct elements $C_n = \{x_1, ..., x_n\} \subset \mathcal{X}$, denote

$$\mathcal{H}_{C_n} = \left\{ (h(x_1), \dots, h(x_n)) : h \in \mathcal{H} \right\},\$$

and therefore $|\mathcal{H}_{C_n}|$ is the number of distinct vectors of length n that can be realized by functions in 451 \mathcal{H} .

452 The n^{th} shatter coefficient is defined as

$$S(\mathcal{H}, n) := \max_{C_n} |\mathcal{H}_{C_n}|.$$

We say that a set C_n is shattered by \mathcal{H} if there exists $f, g : C_n \to \mathcal{Y}$ such that for every $x \in C_n$, 454 $f(x) \neq g(x)$, and

$$\mathcal{H}_C \supseteq \{f(x_1), g(x_1)\} \times \{f(x_2), g(x_2)\} \times \dots \times \{f(x_n), g(x_n)\}$$

If $\mathcal{Y} = \{1, 2\}$, this definition reduces the binary notion of shattering which says all labeling of points can be realized by some function in the hypothesis class \mathcal{H} , i.e., $\mathcal{H}_C = \{1, 2\}^{|C|}$. Contrary to the intuition that multiclass shattering means being able to label all K possible labels for each point $x \in C$: $\mathcal{H}_C = \{1, 2, ..., K\}^{|C|}$. This definition is more like "embed the binary cube into multiclass", where every $x \in C$ is allowed to pick from two labels rather than from K labels.

Definition 3 (Natarajan Dimension) The Natarajan dimension of \mathcal{H} , denoted Ndim(\mathcal{H}), is the maximal size of a shattered set $C \in \mathcal{X}$.

Theorem (Excess Risk Upper Bound of NI-ERM) Let $\epsilon \in [0,1], \kappa \in (0,+\infty)$. Consider any ($P_X, \eta, \tilde{\eta}$) $\in \Pi(\epsilon, \kappa)$, assume function class \mathcal{F} has Natarajan dimension V, and the noisy Bayes classifier \tilde{f}^* belongs to \mathcal{F} . Let $\hat{f} \in \mathcal{F}$ be the ERM trained on $Z^n = \{(X_i, \tilde{Y}_i)\}_{i=1}^n$, then

$$\mathbb{E}_{Z^n} \left[R\left(\hat{f}\right) - R(f^*) \right] \le \epsilon + \frac{1}{\kappa} \cdot 8\sqrt{\frac{V\log 2n + 2V\log K + 4}{2n}} \le \epsilon + \mathcal{O}\left(\frac{1}{\kappa}\sqrt{\frac{V}{n}}\right) \quad up \text{ to log factor.}$$

465 *Proof.* Following directly from Lemma 1, with $(P_X, \eta, \tilde{\eta}) \in \Pi(\epsilon, \kappa)$, we already have

$$R(f) - R(f^*) \le P_X \left(\mathcal{X} \setminus \mathcal{A}_{\kappa} \left(\boldsymbol{\eta}, \widetilde{\boldsymbol{\eta}} \right) \right) + \frac{1}{\kappa} \left(\widetilde{R}(f) - \widetilde{R} \left(\widetilde{f}^* \right) \right)$$
$$\le \epsilon + \frac{1}{\kappa} \left(\widetilde{R}(f) - \widetilde{R} \left(\widetilde{f}^* \right) \right).$$

⁴⁶⁶ Now replace f with NI-ERM \hat{f} , the following procedure is similar to the derivation of generalization

467 error based on VC dimension in binary classification setup, except using a "Natarajan version" of
 468 Sauer's lemma.

Let's start from excess risk bound based on Shattering coefficient.

Lemma 2

$$\mathbb{E}_{Z^n}\left[\widetilde{R}\left(\widehat{f}\right) - \widetilde{R}\left(\widetilde{f^*}\right)\right] \le 8\sqrt{\frac{\log(8eS(\mathcal{H},2n))}{2n}}$$

470 *Proof Sketch.* First convert excess risk $\widetilde{R}(\widehat{f}) - \widetilde{R}(\widetilde{f}^*)$ to generalization error with the inequality 471 ([Devroye et al., 1996, Lemma 8.2])

$$\widetilde{R}\left(\widehat{f}\right) - \widetilde{R}\left(\widetilde{f}^*\right) \le 2 \sup_{f \in \mathcal{F}} \left|\widetilde{R}(f) - \widetilde{R}_n(f)\right|,$$

- where \widetilde{R}_n denotes the empirical risk evaluated on training sample Z^n .
- Then use the "symmetrization lemma" [Vapnik and Chervonenkis, 1971, Lemma 2], for $n\epsilon^2 \ge 2$,

$$\mathbb{P}\left(\sup_{f\in\mathcal{F}}\left|\widetilde{R}(f)-\widetilde{R}_{n}\left(f\right)\right|\geq\epsilon\right)\leq2\mathbb{P}\left(\sup_{f\in\mathcal{F}}\left|\widetilde{R}_{n}'(f)-\widetilde{R}_{n}\left(f\right)\right|\geq\frac{\epsilon}{2}\right)$$

- to convert the population risk $\widetilde{R}(f)$ into $\widetilde{R}'_n(f)$, the risk evaluated on the "ghost sample", a second, independent sample of *n* points.
- Finally, with union bound and Hoeffding's inequality, we can bound the error probability with shattering coefficient on 2n points

$$2\mathbb{P}\left(\sup_{f\in\mathcal{F}}\left|\widetilde{R}'_{n}(f)-\widetilde{R}_{n}(f)\right|\geq\frac{\epsilon}{2}\right)\leq 2S(\mathcal{H},2n)\mathbb{P}\left(\left|\widetilde{R}'_{n}(f)-\widetilde{R}_{n}(f)\right|\geq\frac{\epsilon}{2}\right)$$
$$\leq 2S(\mathcal{H},2n)\left\{\mathbb{P}\left(\left|\widetilde{R}'_{n}(f)-\widetilde{R}(f)\right|\geq\frac{\epsilon}{4}\right)+\mathbb{P}\left(\left|\widetilde{R}'_{n}(f)-\widetilde{R}(f)\right|\geq\frac{\epsilon}{4}\right)\right\}$$
$$\leq 2S(\mathcal{H},2n)\left\{2\cdot2\cdot e^{-2n\left(\frac{\epsilon}{4}\right)^{2}}\right\}$$
$$\leq 8S(\mathcal{H},2n)e^{-\frac{n\epsilon^{2}}{8}},$$

- notice that the inequality also holds for $n\epsilon^2 < 2$ because in this case the upper bound is bigger than 1.
- Therefore we get the convergence in probability bound for ERM \hat{f} ,

$$\mathbb{P}\left(\widetilde{R}\left(\widehat{f}\right) - \widetilde{R}\left(\widetilde{f}^*\right) \ge \epsilon\right) \le 8S(\mathcal{H}, 2n)e^{-\frac{n\epsilon^2}{32}}.$$

480 Convert it to expectation using the inequality from Devroye et al. [1996, Problem 12.1], we finally get

$$\mathbb{E}_{Z^n}\left[\widetilde{R}\left(\widehat{f}\right) - \widetilde{R}\left(\widetilde{f}^*\right)\right] \le 8\sqrt{\frac{\log(8eS(\mathcal{H},2n))}{2n}}.$$

481

The last step is to bound multiclass shattering coefficient with Natarajan dimension. We leverage the following lemma.

Lemma 3 (Natarajan [1989]) Let C and \mathcal{Y} be two finite sets and let \mathcal{H} be a set of functions from C to \mathcal{Y} .

$$|\mathcal{H}| \leq |C|^{Ndim(\mathcal{H})} \cdot |\mathcal{Y}|^{2Ndim(\mathcal{H})}.$$

486 Replacing shattering coefficient with Natarajan dimension, we get

$$\mathbb{E}_{Z^n} \left[\widetilde{R}\left(\widehat{f} \right) - \widetilde{R}\left(\widetilde{f}^* \right) \right] \le 8\sqrt{\frac{\log 8e + \log\left((2n)^V K^{2V} \right)}{2n}} \\ \le 8\sqrt{\frac{V \log 2n + 2V \log K + 4}{2n}}$$

487 Putting things together,

$$\mathbb{E}_{Z^n}\left[R\left(\hat{f}\right) - R(f^*)\right] \le \epsilon + \frac{1}{\kappa} \cdot 8\sqrt{\frac{V\log 2n + 2V\log K + 4}{2n}}$$

488

489 A.2 Experimental Details

490 A.2.1 2D Gaussian with synthetic label noise

For 2D Gaussian mixture data, we draw from two Gaussian centered at $[1 \ 1]^{\top}$ and $[-1 \ -1]^{\top}$, with covariance matrix being identity, 200 data points from each, with label Y = 1, 2 respectively. To generate noisy labels, we flip every label uniformly with some probability. We use Sklearn's logistic regression (with no ℓ_2 regularization). The experiment was conducted on AMD Ryzen 5 3600 CPU. The goal of the simulation is to experimentally verify noise immunity results in Section **??**. Notice that different trial corresponds to different draw of both instances and noisy labels.

Table 2: Testing accuracy of logistic regression on gaussian mixture data with uniform label noise. "Noise rate" refers to $\mathbb{P}(\tilde{Y} \neq Y)$, the percentage of wrong labels in the training data. As theory in Section ?? predicts, when $\mathbb{P}(\tilde{Y} \neq Y)$ reach 50%, there is a sharp decrease in performance.

Noise rate	0	0.05	0.1	0.15	0.2	0.25	0.3	0.35	0.4	0.45	0.5	0.55	0.6	0.65	0.7	0.75	0.8	0.85	0.9	0.95	1
Trial #1	93.00	92.83	92.38	92.08	91.78	91.93	92.25	92.90	91.83	92.58	74.68	25.12	9.70	7.73	7.52	7.25	7.38	7.15	7.18	7.10	7.00
Trial #2	91.73	91.60	92.05	91.63	91.78	91.78	91.68	91.63	91.55	91.48	80.40	21.10	9.93	8.55	8.38	8.22	8.20	8.35	8.33	8.40	8.28
Trial #3	92.73	92.75	92.78	92.78	92.58	92.45	91.68	88.15	82.58	59.83	49.53	35.80	21.28	14.35	9.33	8.53	8.12	7.70	7.13	7.23	7.28
Trial #4	91.55	91.58	91.60	91.63	91.68	91.60	91.25	90.98	89.98	86.38	60.53	9.95	8.75	10.00	10.45	9.08	9.00	9.53	9.20	9.03	8.45
Trial #5	91.55	91.58	91.60	91.63	91.68	91.60	91.25	90.98	89.98	86.38	60.53	9.95	8.75	10.00	10.45	9.08	9.00	9.53	9.20	9.03	8.45
Mean	92.11	92.07	92.08	91.95	91.90	91.87	91.62	90.93	89.18	83.33	65.13	20.40	11.68	10.10	9.23	8.43	8.34	8.45	8.21	8.16	7.89
Std	0.70	0.66	0.51	0.50	0.38	0.35	0.41	1.74	3.79	13.44	12.35	10.94	5.39	2.56	1.29	0.75	0.68	1.07	1.03	0.94	0.70

497 A.2.2 MNIST with synthetic label noise

We flip the clean training label of MNIST (http://yann.lecun.com/exdb/mnist/) uniformly (to any of the wrong classes). We use a shallow neural network with two convolution layers and two fully connected layers. We train with stochastic gradient descent with learning rate 0.01 for 10 epochs, batch size equals 64. We use the same hyperparamters for all tests. The experiments were conducted on a single NVIDIA GTX 1660S GPU. The goal of the simulation is to experimentally verify noise immunity results in Section **??**. Here randomness corresponds to different realization of noisy labels and stochastic gradient descent.

Table 3: Testing accuracy of a shallow CNN (2 conv layers with 2 fully connected layers) on MNIST with uniform label noise. "Noise rate" refers to $\mathbb{P}(\tilde{Y} \neq Y)$, the percentage of wrong labels in the training data. As theory in Section ?? predicts, when $\mathbb{P}(\tilde{Y} \neq Y)$ reach 90%, there is a sharp decrease in performance.

Noise rate	0	0.05	0.1	0.15	0.2	0.25	0.3	0.35	0.4	0.45	0.5	0.55	0.6	0.65	0.7	0.75	0.8	0.85	0.9	0.95	1
Trial #1	98.97	98.89	98.81	98.46	98.49	98.16	98.46	98.07	97.98	97.57	97.88	97.84	97.19	97.10	96.70	95.02	89.00	83.72	11.58	0.17	0.03
Trial #2	98.88	98.73	98.94	98.55	98.72	98.66	98.50	98.24	98.15	98.23	97.86	97.98	97.70	97.10	96.91	95.76	91.99	88.49	9.99	0.08	0.04
Trial #3	99.00	99.04	98.86	98.56	98.69	98.66	98.51	98.49	98.37	98.25	98.25	97.39	97.37	97.18	96.66	94.88	92.15	81.48	6.19	0.14	0.04
Trial #4	99.04	98.86	98.70	98.76	98.83	98.65	98.34	98.42	98.58	98.47	98.00	97.41	97.63	97.09	96.46	95.94	93.19	84.78	8.68	0.19	0.01
Trial #5	99.05	98.58	98.89	98.82	98.72	98.83	98.34	98.55	98.40	98.38	98.01	97.31	97.33	96.21	96.29	94.92	90.38	85.84	8.98	0.13	0.08
Mean	98.99	98.82	98.84	98.63	98.69	98.59	98.43	98.35	98.30	98.18	98.00	97.59	97.44	96.94	96.60	95.30	91.34	84.86	9.08	0.14	0.04
Std	0.07	0.17	0.09	0.15	0.12	0.25	0.08	0.20	0.23	0.36	0.16	0.30	0.21	0.41	0.24	0.51	1.65	2.59	1.98	0.04	0.03

505 A.2.3 CIFAR with synthetic label noise

We flip the clean training label of CIFAR-10 (https://www.cs.toronto.edu/~kriz/cifar. 506 html) uniformly (to any of the wrong classes). To have a fair comparison between different methods, 507 we fix the realization of noisy labels. Follow the 2-step procedure described in Section 5, we use 508 different pre-trained neural networks as feature extractor: forward-passing the training image thought 509 the network and record the feature. Then use sklearn's (https://scikit-learn.org/stable/) 510 logistic regression function to fit the (feature, noisy label) pair in a full batch manner. We pre-511 specify a range of values for ℓ_2 regularization ($\{0.0001, 0.001, 0.01, 0.1, 1, 10, 100\}$) and number of 512 iterations for lbfgs optimizer ($\{10, 20, 50, 100\}$), then do cross-validation on noisy data to pick the 513 best hyper-parameters. We use the same range of hyper-parameters in all tests. The experiments were 514 conducted on a single NVIDIA Tesla V100 GPU. The result is deterministic. 515

Table 4: Peformance on CIFAR-10 with synthetic label noise. We apply linear model on top of different feature extractors: "ResNet-50 TL" refers to using a pre-trained ResNet-50 on ImageNet [Deng et al., 2009] (available in Pytorch model library) in a transfer learning fashion, "ResNet-50 SSL" refers to using a pre-trained ResNet-50 on unlabeled CIFAR data with self-supervised loss [Chen et al., 2020] (publicly downloadable weights https://github.com/ContrastToDivide/C2D? tab=readme-ov-file) and "DINOv2 SSL" refers to using the self-supervised foundation model DINOv2 [Oquab et al., 2023] (available at https://github.com/facebookresearch/dinov2) as the feature extractor. "Noise rate" refers to $\mathbb{P}(\tilde{Y} \neq Y)$, the percentage of wrong labels in the training data. As theory in Section ?? predicts, when $\mathbb{P}(\tilde{Y} \neq Y)$ reach 90%, there is a sharp decrease in performance. We employed Python's sklearn logistic regression and cross-validation functions without data augmentation. The results are deterministic.

Noise rate	0	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.85	0.9	0.95	1
Linear	41.37	41.09	40.97	40.37	40.45	39.44	37.28	35.20	26.74	18.00	10.28	5.50	3.92
Linear + ResNet-50 TL	90.17	89.58	89.01	88.27	87.55	87.28	86.40	85.01	82.03	74.02	10.82	1.47	0.26
Linear + ResNet-50 SSL	92.48	92.26	91.74	91.46	91.13	90.33	91.07	90.99	89.11	83.89	10.08	1.31	0.34
Linear + DINOv2 SSL	99.25	99.27	99.23	99.14	99.10	99.11	99.02	98.84	95.50	76.91	10.13	0.92	0.03

516 A.2.4 CIFAR with human label error

⁵¹⁷ We load the noisy human labels provided by http://noisylabels.com/, then follow exact the same procedure as above.

Table 5: Performance on CIFAR-N dataset (http://noisylabels.com/) in terms of testing accuracy. "Aggre", "Rand1", ..., "Noisy" denote various types of human label noise. We apply linear model on top of different feature extractors: "ResNet-50 TL" refers to using a pre-trained ResNet-50 on ImageNet [Deng et al., 2009] in a transfer learning fashion, "ResNet-50 SSL" refers to using a pre-trained ResNet-50 on unlabeled CIFAR data with self-supervised loss [Chen et al., 2020] and "DINOv2 SSL" refers to using the self-supervised foundation model DINOv2 [Oquab et al., 2023] as the feature extractor. We employed Python's sklearn logistic regression and cross-validation functions without data augmentation; the results are deterministic and directly reproducible.

Methods		CIFAR-10N								
	Aggre	Rand1	Rand2	Rand3	Worst	Noisy				
Linear	40.73	40.41	40.31	40.63	38.43	16.61				
Linear + ResNet-50 TL	89.18	88.63	88.61	88.66	85.32	62.89				
Linear + ResNet-50 SSL	91.78	91.66	91.39	91.28	87.84	57.95				
Linear + DINOv2 SSL	98.69	98.80	98.65	98.67	95.71	83.17				

519 A.2.5 Linear probing, then fine tuning (LP-FT)

520 We study whether LP-FT works in label noise learning scenario.

Table 6: Performance on CIFAR-N dataset (http://noisylabels.com/) in terms of testing accuracy. "Aggre", "Rand1", ..., "Noisy" denote various types of human label noise. We apply linear model on top of different feature extractors: "ResNet-50 TL" refers to using a pre-trained ResNet-50 on ImageNet [Deng et al., 2009] in a transfer learning fashion, "ResNet-50 SSL" refers to using a pre-trained ResNet-50 on unlabeled CIFAR data with contrastive loss [Chen et al., 2020] and "DINOv2 SSL" refers to using the self-supervised foundation model DINOv2 [Oquab et al., 2023] as the feature extractor.

Feature	Method			CIFA	R-10N			CIFAR-100N		
		Clean	Aggre	Rand1	Rand2	Rand3	Worst	Clean	Noisy	
D. N. (50 TI	LP	90.17	89.18	88.63	88.61	88.66	85.32	71.79	62.89	
ResNet-50 1L	LP-FT	95.94	92.03	88.55	87.78	87.82	71.88	82.3	63.85	
DecNet 50 CCI	LP	92.54	91.78	91.66	91.46	91.17	87.85	69.88	57.98	
Keshel-30 SSL	LP-FT	94.11	89.11	84.49	83.75	84.15	65.00	74.41	54.49	
DINOv2 (small) SSL	LP LP-FT	96.09 98.23	94.8 93.29	94.39 88.03	94.42 87.27	94.35 86.94	91.14 67.42	83.82 89.97	72.46 64.81	

521 A.2.6 Robust learning strategy over DINOv2 feature

This section examines how different robust learning strategy works over DINOv2 feature, compared with only training with cross entropy.

	Feature	Method			CIFA	R-10N			CIFAR	CIFAR-100N		
			Clean	Aggre	Rand1	Rand2	Rand3	Worst	Clean	Noisy		
524	DINOv2 SSL	CE MAE Sigmoid CE + SAM	99.25 99.27 99.26 99.09	98.69 99.04 98.86 97.66	98.8 99.01 98.91 98.47	98.65 99.09 98.87 98.53	98.67 99.11 98.96 98.47	95.71 95.55 96.66 95.47	92.85 90.68 92.82 89.97	83.17 82.55 82.03 82.85		

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