STATISTICAL TEST FOR ANOMALY DETECTIONS US ING VARIATIONAL AUTO-ENCODERS BY SELECTIVE INFERENCE

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Abstract

Over the past decade, Variational Autoencoders (VAE) have become a widely used tool for anomaly detection (AD), with research advancing from algorithm development to real-world applications. However, a critical challenge remains—the lack of a reliable method to rigorously assess the reliability of detected anomalies, which restricts its use in high-stakes decision-making tasks such as medical diagnostics. To overcome this limitation, we introduce the VAE-AD Test, a novel approach for quantifying the statistical reliability of VAE-based AD. The key advantage of the VAE-AD Test lies in its ability to properly control the probability of misidentifying anomalies under a pre-specified level of guarantee α (e.g., 0.05). Specifically, by carefully analyzing the AD process of VAE, which operates through piecewiselinear functions, and leveraging the Selective Inference (SI) framework to assign valid p-values to the detected anomalies, we prove that theoretical control of the false detection rate is achievable. Experiments conducted on both synthetic and real-world datasets robustly support our theoretical results, showcasing the VAE-AD Test's superior performance. To our knowledge, this is the first work capable of conducting valid statistical inference to assess the reliability of VAE-based AD.

1 INTRODUCTION

031 Anomaly detection (AD) is the process of identifying unusual deviations in data that do not conform 032 to expected behavior. AD is crucial across various domains because it provides early warnings of 033 potential issues, thereby enabling timely interventions to prevent critical events. Traditional AD 034 techniques, while effective in simple scenarios, frequently fall short when dealing with complex data, thus motivating the use of deep learning-based AD to better handle such complexities. In this study, we focus on AD using the Variational Auto-Encoder (VAE), and its application to medical images. In the training phase of VAE-based AD, the VAE learns the distribution of normal images 037 by training exclusively on images that do not contain abnormal regions. The parameters of a VAE are optimized to minimize the reconstruction error, thereby learning a compressed representation of the normal data. In the test phase, when a test image is fed into the trained VAE, the model 040 attempts to reconstruct the image based on its learned representation. Since the VAE is trained on 041 normal data, it would successfully reconstruct the normal regions of the image, while it would fail 042 to properly reconstruct the abnormal regions that were not included in the normal data. Therefore, 043 regions with large reconstruction errors are detected as abnormal regions. Figure 1 shows an example 044 of VAE-based AD for brain tumor images.

When VAE-based AD is employed for high-stakes decision-making tasks, such as medical diagnosis, there is a significant risk that model inaccuracies might lead to critical errors, potentially resulting in false detections. To address this issue, we develop a statistical test for VAE-based AD, which we call *VAE-AD Test*. The proposed VAE-AD test enables us to obtain a quantifiable and interpretable measure for the detected anomaly region in the form of *p*-value. The obtained *p*-value represents the probability that the detected anomaly regions are obtained by chance due to the randomness contained in the data. It is important to note that the statistical test for detected abnormal regions is considered as a *data-driven hypothesis*, as the abnormal region is selected based on the test image itself. In other words, since both of the selection of the hypothesis (selection of abnormal regions) and the evaluation of the hypothesis (evaluation of abnormal regions) are performed on the same data, 054

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(a) Image without tumor region. $p_{\text{naive}} = 0.000$ (false detection) and $p_{\text{selective}} = 0.668$ (true negative).



(b) Image with tumor region. $p_{\text{naive}} = 0.000$ (true detection) and $p_{\text{selective}} = 0.000$ (true detection).

Figure 1: An illustration of the proposed VAE-AD Test in brain image analysis. When an anomaly region is detected based on the difference between the original and the reconstructed images by a VAE, the VAE-AD Test provides a *p*-value to quantify its statistical reliability. The upper plot shows the results of the VAE-AD Test and the conventional method, the latter of which does not consider the fact that the anomaly region is detected by VAE. The lower plot shows the results for a case with anomaly regions. With the proposed method ($p_{\text{selective}}$), correct decisions are made in both cases; the former has a large *p*-value and the latter has a small *p*-value. In contrast, with the conventional method (p_{naive}), both *p*-values are small, indicating false detection in the former case.

applying traditional statistical test to the selected hypothesis leads to selection bias. Therefore, in this
 study, we introduce the *Conditional Selective Inference (CSI)* framework to remove the selection bias.

Related Works. Over the last decade, there has been a significant pursuit in applying deep learning techniques to AD problems (Chalapathy & Chawla, 2019; Pang et al., 2021; Tao et al., 2022). A large number of studies have been conducted for unsupervised AD using VAEs (Baur et al., 2021; Chen & Konukoglu, 2018; Chow et al., 2020; Jana et al., 2022). In this study, we focus on the task of identifying the anomalous regions in the input image, which is called *anomaly localization* within the AD tasks (Zimmerer et al., 2019; Lu & Xu, 2018; Baur et al., 2019).

There are mainly two research directions for improving VAEs for AD. The first direction is on 085 improving the detection rate (Zimmerer et al., 2019; Dehaene et al., 2020), while the second direction is on modifying the VAE itself to make it suitable for AD (Baur et al., 2019; Chen & Konukoglu, 2018; 087 Wang et al., 2020). However, to our knowledge, there has been no existing studies for quantifying the 088 statistical reliability of detected abnormal regions with theoretical validity. In traditional statistical 089 tests, the hypothesis needs to be predetermined and must remain independent of the data. However, 090 in data-driven approaches, it is necessary to select hypotheses based on the data and then assess 091 the reliability of the hypotheses using the same data. This issue, known as *double dipping*, arises 092 because the same data is used for both the selection and evaluation of hypotheses, leading to selection bias (Breiman, 1992). Because anomalies are detected based on data (a test image), when evaluating the reliability of the detected anomalies using the same data, the issue of selection bias arises. 094

CSI has recently gained attention as a framework for statistical hypothesis testing of data-driven hypotheses (Lee et al., 2016; Taylor & Tibshirani, 2015). CSI was initially developed for the statistical inference of feature selection in linear models (Fithian et al., 2015; Tibshirani et al., 2016; Loftus & Taylor, 2014; Suzumura et al., 2017; Le Duy & Takeuchi, 2021; Sugiyama et al., 2021; Duy & Takeuchi, 2022), then extended to various problems (Lee et al., 2015; Choi et al., 2017; Chen & Bien, 2020; Tanizaki et al., 2020; Duy et al., 2020; Gao et al., 2022; Le Duy et al., 2024), and later to neural networks (Duy et al., 2022; Miwa et al., 2023; Shiraishi et al., 2024; Katsuoka et al., 2024), but none of these studies focused on inference on VAE.

Contributions. To our knowledge, this is the first formulation of an approach that provides a quantifiable and interpretable measure for the reliability of VAE-based AD, presented in the form of a *p*-value within a statistical testing framework. The second contribution is the development of an SI method for VAEs, which entails characterizing the hypothesis selection event by a VAE. Finally, our third contribution is demonstrating the effectiveness of the proposed VAE-AD Test through numerical experiments with synthetic data and brain tumor images.

108 2 ANOMALY DETECTION (AD) BY VAE

Variational Autoencoder (VAE). VAEs are generative models consisting of an encoder network 111 and a decoder network Kingma & Welling (2013). Given an input image (denoted by $x \in \mathbb{R}^n$), 112 it is encoded as a latent vector (denoted by $z \in \mathbb{R}^m$), and the latent vector is decoded back to 113 the input image, where n is the number of pixels of an image and m is the dimension of a latent vector. In the generative process, it is assumed that a latent vector z is sampled from a prior 114 distribution $p_{\theta^*}(z)$ and then, image x is sampled from a conditional distribution $p_{\theta^*}(x|z)$. The 115 prior distribution $p_{\theta^*}(z)$ and the conditional distribution $p_{\theta^*}(x|z)$ belongs to family of distributions 116 parametrized by θ and θ^* denotes the true value of the parameter. The encoder network approximates 117 the posterior distribution $p_{\theta}(z|x)$ by the parametric distribution $q_{\phi}(z|x)$, where ϕ represents the 118 set of parameters, while the decoder network estimates the conditional distribution by $p_{\theta}(x|z)$. 119 The encoder and the decoder networks of a VAE are trained by maximizing so-called evidence 120 lower bound (ELBO): $L_{\theta,\phi} = \mathbb{E}_{q_{\theta}(\boldsymbol{z}|\boldsymbol{x})} \left[\log p_{\phi}(\boldsymbol{x}|\boldsymbol{z}) \right] - \mathrm{KL} \left[q_{\theta}(\boldsymbol{z}|\boldsymbol{x}) || p(\boldsymbol{z}) \right]$, where $\mathrm{KL} \left[\cdot || \cdot \right]$ is the 121 Kullback-Leibler divergence between two distributions. We model the approximated posterior 122 distribution $q_{\phi}(z|x)$ as a normal distribution $N(\mu_{\phi}(x), I_n \sigma_{\phi}^2(x))$, where $\mu_{\phi}(x)$ and $\sigma_{\phi}^2(x)$ are the 123 outputs of the encoder network. The conditional distribution $p_{\theta}(\boldsymbol{x}|\boldsymbol{z})$ is also modeled as a normal 124 distribution $N(\mu_{\theta}(z), I_n)$, where $\mu_{\theta}(z)$ is the output of the decoder network. Furthermore, the prior 125 distribution $p_{\theta^*}(z)$ is modeld as a standard normal distribution $N(0, I_m)$. The structure of the VAE 126 used in this study is shown in Appendix A.1.

Anomaly Detection Using VAEs. VAEs can be effectively used for anomaly localization task. The 128 goal of anomaly localization is to identify the abnormal region within a given test image. In the 129 training phase, we assume that only normal images (e.g., brain images without tumors) are available. 130 A VAE is trained on normal images to learn a compact representation of the normal image distribution 131 in the latent space. In the test phase, a test image x is fed into the trained VAE, and a reconstructed 132 image is obtained by using the encoder and the decoder as $\hat{x} = \mu_{\theta} (\mu_{\phi}(x))$. Since the VAE is 133 trained only on normal images, normal region in the test image would be reconstructed well, whereas 134 the reconstruction error of abnormal regions would be high. Therefore, it is reasonable to define the 135 degree of anomaly of each pixel as 136

$$E_i(\boldsymbol{x}) = |x_i - \hat{x}_i|, i \in [n], \tag{1}$$

where x_i and \hat{x}_i is the *i*th pixel value of x and \hat{x} , respectively. Using a user-specified threshold 138 $\lambda > 0$, the anomaly region of a test image x is defined as 139

> $A_{\boldsymbol{x}} = \{ i \in |n| \mid E_i(\boldsymbol{x}) \ge \lambda \}.$ (2)

141 As for the definition of the anomaly region, there are possibilities other than those given by Eqs. (1) 142 and (2). In this paper, we proceed with these choices, but the proposed VAE-AD Test is generally 143 applicable to other choices.

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3 STATISTICAL TEST FOR ABNORMAL REGIONS

147 Statistical model of an image. To formulate the reliability assessment of the abnormal region as a 148 statistical testing problem, it is necessary to introduce a statistical model of an image. In this study, 149 an image is considered as a sum of true signal component $s \in \mathbb{R}^n$ and noise component $\epsilon \in \mathbb{R}^n$. 150 Regarding the true signal component, each pixel can have an arbitrary true signal value without any 151 particular assumption or constraint. On the other hand, regarding the noise component, it is assumed 152 to follow a normal distribution, and their covariance matrix is estimated using normal data different from that used for the training of the VAE. Namely, an image with n pixels can be represented as an 153 n-dimensional random vector 154

$$\boldsymbol{X} = (X_1, \dots, X_n) = \boldsymbol{s} + \boldsymbol{\epsilon}, \ \boldsymbol{\epsilon} \sim N(\boldsymbol{0}, \boldsymbol{\Sigma}), \tag{3}$$

156 where $s \in \mathbb{R}^n$ is the true signal vectors, and $\epsilon \in \mathbb{R}^n$ is the noise vector with covariance matrix Σ . In 157 the following, the capital X denotes an image as a random vector, while the lowercase x represents 158 an observed image. To formulate the statistical test, we consider the AD using VAEs in Eq. equation 2 159 as a function \mathcal{A} that maps a random input image X to the abnormal region A_X , i.e.,

$$\mathcal{A}: \mathbb{R}^n \ni \mathbf{X} \mapsto A_{\mathbf{X}} \in 2^{[n]},\tag{4}$$

where $2^{[n]}$ is the power set of $[n] := \{1, 2, ..., n\}$.

Formulation of statistical test. Our goal is to make a judgment whether the abnormal region A_X merely appears abnormal due to the influence of random noise, or if there is a true anomaly in the true signal in the abnormal region. In order to quantify the reliability of the detected abnormal region, the statistical test is performed for the difference between the true signal in the abnormal region $\{s_i\}_{i \in A_X}$ and the true signal in the normal region $\{s_i\}_{i \in A_X^c}$ where A_X^c is the complement of the abnormal region. In this study, as an example, we consider the hypothesis for the difference in true mean signals between A_X and A_X^c by considering the following null and alternative hypotheses:

$$H_{0}: \frac{1}{|A_{\mathbf{X}}|} \sum_{i \in A_{\mathbf{X}}} s_{i} = \frac{1}{|A_{\mathbf{X}}^{c}|} \sum_{i \in A_{\mathbf{X}}^{c}} s_{i}, \text{ v.s. } H_{1}: \frac{1}{|A_{\mathbf{X}}|} \sum_{i \in A_{\mathbf{X}}} s_{i} \neq \frac{1}{|A_{\mathbf{X}}^{c}|} \sum_{i \in A_{\mathbf{X}}^{c}} s_{i}.$$
(5)

For clarity, we mainly consider a test for the mean difference as a specific example — however, the proposed VAE-AD Test is applicable to a more general class of statistical tests. Specifically, let $\eta \in \mathbb{R}^n$ be an arbitrary *n*-dimensional vector depending on the abnormal region A_X . Then, the proposed method can cover a statistical test represented as

$$\mathbf{H}_0: \boldsymbol{\eta}^\top \boldsymbol{s} = c \quad \text{v.s.} \quad \mathbf{H}_1: \boldsymbol{\eta}^\top \boldsymbol{s} \neq c, \tag{6}$$

where c is an arbitrary constant. The formulation in Eq. (6) covers a wide range of practically useful statistical tests. In fact, Eq. (5) is a special case of Eq. (6). It can cover differences not only in means but also in other measures such as maximum difference, and differences after applying some image filters (e.g., Gaussian filter).

Test statistic. To evaluate the hypothesis defined in Eq. equation 5, we define the test statistic as

$$T(\boldsymbol{X}) = \frac{1}{|A_{\boldsymbol{X}}|} \sum_{i \in A_{\boldsymbol{X}}} X_i - \frac{1}{|A_{\boldsymbol{X}}^c|} \sum_{i \in A_{\boldsymbol{X}}^c} X_i = \boldsymbol{\eta}^\top \boldsymbol{X},$$
(7)

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where $\boldsymbol{\eta} = \frac{1}{|A_{\boldsymbol{X}}|} \mathbf{1}_{A_{\boldsymbol{X}}} - \frac{1}{|A_{\boldsymbol{X}}^c|} \mathbf{1}_{A_{\boldsymbol{X}}^c}$ and $\mathbf{1}_A \in \mathbb{R}^n$ is a vector with 1 if $i \in A_{\boldsymbol{X}}$ and 0 otherwise.

Naive *p***-values.** When the test statistic in Eq. (7) is used for the statistical test in Eq. (5), the *p*-value 188 can be easily calculated if η does not depend on the image X, i.e., if the abnormal region A_X is 189 detected without looking at the X. In this unrealistic situation, the p-value, which we call naive 190 *p-value* can be computed as $p_{\text{naive}} = \mathbb{P}_{H_0}(|T(\boldsymbol{X})| \ge |T(\boldsymbol{x})|)$, where \boldsymbol{X} is a random vector and \boldsymbol{x} is 191 the observed image. Under the unrealistic assumption, the p_{naive} can be easily computed because 192 the null distribution of $T(\mathbf{X}) = \boldsymbol{\eta}^{\top} \mathbf{X}$ is normally distributed with $N(0, \boldsymbol{\eta}^{\top} \Sigma \boldsymbol{\eta})$. Unfortunately, 193 however, in the actual situation where η depends on X, a statistical test using p_{naive} is *invalid* in the 194 sense that $P_{H_0}(p_{\text{naive}} \leq \alpha) > \alpha$, $\exists \alpha \in [0, 1]$. Namely, the probability of Type I error (an error that 195 a normal region is mistakenly detected as anomaly) cannot be controlled at the desired level α . 196

4 CONDITIONAL SELECTIVE INFERENCE (CSI) FOR VAE-BASED AD

In this section, we present the proposed VAE-AD Test, a valid statistical test for VAE-based AD task.

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4.1 CONDTIONAL SELECTIVE INFERENCE (CSI)

In CSI, *p*-values are computed based on the null distribution conditional on a event that a certain hypothesis is selected. The goal of CSI is to compute a *p*-value that satisfies
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$$P_{\mathrm{H}_{0}}(p \le \alpha \mid A_{\boldsymbol{X}} = A) \le \alpha, \tag{8}$$

where the condition part $A_{\mathbf{X}} = A$ in Eq. (8) indicates that we only consider images \mathbf{X} for which a certain hypothesis (abnormal region) A is detected. If the conditional type I error can be controlled as in Eq. (8) for all possible hypotheses $A \in 2^{[n]}$, then, by the law of total probability, the marginal type I error can also be controlled for all $\alpha \in (0, 1)$ because

$$P_{\mathcal{H}_0}(p \le \alpha) = \sum_{A \in 2^{[n]}} P_{\mathcal{H}_0}(A)(p \le \alpha \mid A_{\mathbf{X}} = A) \le \alpha.$$

Therefore, in order to perform valid statistical test, we can employ *p*-values conditional on the hypothesis selection event. To compute a *p*-value that satisfies Eq. (8), we need to derive the sampling distribution of the test-statistic

$$T(\boldsymbol{X})|\{A_{\boldsymbol{X}} = A_{\boldsymbol{x}}\}.$$
(9)

4.2 CSI FOR PIECEWISE-ASSIGNMENT FUNCTIONS 217

We derive the CSI for algorithms expressed in the form of a *piecewise-assignment function*. Later on, we show that the mapping $\mathcal{A} : \mathbf{X} \mapsto A_{\mathbf{X}}$ in Eq. (4) is a piecewise-assignment function, and this will result in the proposed VAE-AD Test.

Definition 1 (Piecewise-Assignment Function). Let us consider a function $M : \mathbb{R}^n \ni X \mapsto M_X \in \mathcal{M}$ which assigns an image X to a hypothesis among a finite set of hypotheses \mathcal{M} . We call the function M a piecewise-assignment function if it is written as

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253 254 $M_{\boldsymbol{X}} = \begin{cases} M_1, & \text{if } \boldsymbol{X} \in \mathcal{P}_1^M, \\ \vdots & & \\ M_k, & \text{if } \boldsymbol{X} \in \mathcal{P}_k^M, \\ \vdots & & \\ M_{K^M}, & \text{if } \boldsymbol{X} \in \mathcal{P}_{K^M}^M, \end{cases}$ (10)

where \mathcal{P}_k^M , $k \in [K^M]$, represents a polytope in \mathbb{R}^n which can be written as $\mathcal{P}_k^M = \{ X \in \mathbb{R}^n \mid \Delta_k^M X' \leq \delta_k^M \}$ using a certain matrix Δ_k^M and a vector δ_k^M with appropriate sizes, and K^M is the number of polytopes. Here, we note that the same hypothesis may be assigned to different polytopes.

When a hypothesis is selected by a piecewise-assignment function in the form of Eq. (10), the following theorem tells that the conditional *p*-value that satisfies Eq. (8) can be derived by using truncated normal distribution.

Theorem 1. Consider a random image X and an observed image x. Let M_X and M_x be the hypotheses obtained by applying a piecewise-assignment function in the form of Eq. (10) to X and x, respectively. Let $\eta \in \mathbb{R}^n$ be a vector depending on M_x , and consider a test statistic in the form of $T(X) = \eta^\top X$. Furthermore, define

$$\mathcal{Q}_{\boldsymbol{X}} = \left(I_n - \frac{\Sigma \eta \eta^{\top}}{\eta^{\top} \Sigma \eta}\right) \boldsymbol{X} \text{ and } \mathcal{Q}_{\boldsymbol{x}} = \left(I_n - \frac{\Sigma \eta \eta^{\top}}{\eta^{\top} \Sigma \eta}\right) \boldsymbol{x}$$

246 Then, the conditional distribution

$$T(\boldsymbol{X}) \mid \{M_{\boldsymbol{X}} = M_{\boldsymbol{x}}, \mathcal{Q}_{\boldsymbol{X}} = \mathcal{Q}_{\boldsymbol{x}}\}$$

is a truncated normal distribution $TN(\eta^{\top}\mu, \eta^{\top}\Sigma\eta; Z)$ with the mean $\eta^{\top}\mu$, the variance $\eta^{\top}\Sigma\eta$, and the truncation intervals Z. The truncation intervals Z is represented as

$$\mathcal{Z} = \bigcup_{k:M_k = M_{\pi}} [L_k^M, U_k^M],$$

where, for $k \in [K^M]$, L_k^M and U_k^M are defined as follows:

$$L_k^M = \max_{j:(\boldsymbol{\beta}_k^M)_j > 0} \frac{(\boldsymbol{\alpha}_k^M)_j}{(\boldsymbol{\beta}_k^M)_j}, \ U_k^M = \min_{j:(\boldsymbol{\beta}_k^M)_j < 0} \frac{(\boldsymbol{\alpha}_k^M)_j}{(\boldsymbol{\beta}_k^M)_j}$$

with
$$\boldsymbol{\alpha}_k^M = \boldsymbol{\delta}_k^M - \boldsymbol{\Delta}_k^M \mathcal{Q}_{\boldsymbol{x}}$$
 and $\boldsymbol{\beta}_k^M = \boldsymbol{\Delta}_k^M \Sigma \boldsymbol{\eta} (\Sigma \boldsymbol{\eta}^\top \Sigma \boldsymbol{\eta})^{-1}$

The proof of Theorem 1 is deferred to Appendix A.2. Using the sampling distribution of the test statistic $T(\mathbf{X})$ conditional on $\{M_{\mathbf{X}} = M_{\mathbf{x}}, Q_{\mathbf{X}} = Q_{\mathbf{x}}\}$ in Theorem 1, we can define the *p*-value as

$$p_{\text{selective}} = \mathbb{P}_{H_0}(|T(\boldsymbol{X})| \ge |T(\boldsymbol{x})| \mid M_{\boldsymbol{X}} = M_{\boldsymbol{x}}, \mathcal{Q}_{\boldsymbol{X}} = \mathcal{Q}_{\boldsymbol{x}}).$$
(11)

The selective *p*-value $p_{\text{selective}}$ defined in Eq. equation 11 satisfies

$$\mathbb{P}_{\mathrm{H}_{0}}(p_{\mathrm{selective}} \leq \alpha \mid M_{\boldsymbol{X}} = M_{\boldsymbol{x}}) = \alpha, \; \forall \alpha \in [0, 1]$$

because Q_X is independent of the test statistic $T(X) = \eta^{\top} X$. From the discussion in §4.1, a valid statistical test can be conducted by using $p_{\text{selective}}$ in Eq. (11).

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4.3 PIECEWISE-LINEAR FUNCTIONS

We showed that, if the hypothesis selection algorithm is represented in the form of piecewiseassignment function, we can formulate valid selective *p*-values. The purpose of this subsection is to set the stage for demonstrating in the next subsection how the entire process of a trained VAE can be depicted as a *piecewise-linear function*, and how VAE-based AD algorithm in Eq. (4) is represented as a piecewise-assignment function.

Definition 2 (Piecewise-Linear Function). A piecewise-linear function $f : \mathbb{R}^n \to \mathbb{R}^m$ is written as: ($\mathbf{W}^f \mathbf{V} \mapsto \mathbf{f}^f \mathbf{V} = \mathcal{D}^f$

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 $f(\boldsymbol{X}) = \begin{cases} \Psi_{1}^{f} \boldsymbol{X} + \psi_{1}^{f}, & \text{if } \boldsymbol{X} \in \mathcal{P}_{1}^{f}, \\ \vdots & & \\ \Psi_{k}^{f} \boldsymbol{X} + \psi_{k}^{f}, & \text{if } \boldsymbol{X} \in \mathcal{P}_{k}^{f}, \\ \vdots & & \\ \Psi_{Kf}^{f} \boldsymbol{X} + \psi_{Kf}^{f}, & \text{if } \boldsymbol{X} \in \mathcal{P}_{Kf}^{f}, \end{cases}$ (12)

where \mathcal{P}_k^f represents a polytope in \mathbb{R}^n written as $\mathcal{P}_k^f = \{ \mathbf{X} \in \mathbb{R}^n \mid \mathbf{\Delta}_k^f \mathbf{X}' \leq \mathbf{\delta}_k^f \}$ for $k \in K^f$ with a certain matrix $\mathbf{\Delta}_k^f$ and a vector $\mathbf{\delta}_k^f$ with appropriate sizes. Furthermore, Ψ_k^f and ψ_k^f for $k \in K^f$ are the *k*-th linear transformation matrix and the bias vector, respectively, and K^f denotes the number of polytopes of a piecewise-linear function f.

Considering piecewise-assignment and piecewise-linear functions, the following properties straightforwardly hold:

• The concatenation of two or more piecewise-linear functions results in a piecewise-linear function.

• The composition of two or more piecewise-linear functions results in a piecewise-linear function.

• The composition of a piecewise-linear function and a piecewise-assignment function results in a piecewise-assignment function.

4.4 VAE-BASED AD AS PIECEWISE-ASSIGNMENT FUNCTION

In this subsection, we show that the VAE-based AD algorithm in Eq. (4) is a piecewise-assignment function by verifying that i) the reconstruction error in Eq. (1) is a piecewise-linear function, and ii)
 the thresholding in Eq. (2) is a piecewise-assignment function.

303 Most of basic operations and common activation functions used in the encoder and decoder networks 304 can be represented as piecewise-linear functions in the form of Eq. (12). For example, the ReLU 305 function is a piecewise-linear function. Operations like matrix-vector multiplication, convolution, 306 and upsampling are linear, which categorizes them as special cases of piecewise-linear functions 307 Furthermore, operations like max-pooling and mean-pooling can be represented in the form of Eq. (12). For instance, max-pooling of two variables can be expressed as $\max\{u, v\} = u \cdot I(u \ge u)$ 308 $v) + v \cdot I(v > u)$, which is a piecewise-linear function with $\tilde{K}^{f} = 2$. Consequently, the encoder 309 and decoder networks of the VAE, composed or concatenated from piecewise-linear functions, form 310 a piecewise-linear function. We note that this characteristic is not exclusive to our VAE; instead, it 311 applies to the majority of CNN-type deep learning models¹. 312

Furthermore, the reconstruction error in Eq. (1) is also a piecewise-linear function. Specifically, let f_{abs} be the absolute value function, which is clearly piecewise-linear function, f_{mm1} be a function for multiplying the matrix $(I_n, -I_n)$ from the left, and f_{mm2} be a function for multiplying the matrix $(I_n, I_n)^{T}$ from the left. Then, the reconstruction error $E_i(\mathbf{X}) = |\boldsymbol{\mu}_{\boldsymbol{\theta}}(\boldsymbol{\mu}_{\boldsymbol{\phi}}(\mathbf{X})) - \mathbf{X}|_i$ is given as the *i*th element of the following compositions of multiple piecewise-linear functions:

$$f_{\rm abs} \circ f_{\rm mm1} \circ [\ \boldsymbol{\mu}_{\boldsymbol{\theta}} \circ \boldsymbol{\mu}_{\boldsymbol{\phi}} \quad I_n \] \circ f_{\rm mm2}(\boldsymbol{X}).$$

The thresholding operation in Eq. (2) is clearly piecewise-assignment function. It means that the operation of detecting abnormal region A_X in Eq. (4) is composition of piecewise-linear function

 ¹An example of components that do not exhibit piecewise linearity is nonlinear activation function such as
 the sigmoid function. However, since a one-dimensional nonlinear function can be approximated with high accuracy by a piecewise-linear function with sufficiently many segments, there are no practical problems.

and piecewise-assignment function, which results in a piecewise-assignment function. We summarize the aforementioned discussion into the following lemma.
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Lemma 1. The anomaly detection using VAE defined in Eq. equation 4, which uses piecewise-linear functions in the encoder and decoder network, is a piecewise-assignment function.

Consequently, we can conduct the statistical test in equation 5 based on the selective *p*-value in equation 11 along with Theorem 1.

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5 COMPUTATIONAL TRICKS

334 In this section, we demonstrate the procedure for efficiently computing the truncated intervals \mathcal{Z} 335 derived from Eq. equation 4. The identification of \mathcal{Z} is challenging because the VAE-based AD is 336 comprised of a substantial number of known piecewise-linear functions and a piecewise-assignment 337 function. There are two difficulties: i) which indices of k whose anomaly region is the same 338 as the observed one, and ii) how to compute each truncated interval $[L_k^A, U_k^A]$. Our idea is to leverage parametric programming in conjunction with *auto-conditioning* to efficiently compute Z. 339 Specifically, we can identify only the necessary indices of k and determining their respective intervals 340 $[L_k^a, U_k^a]$. This enables us to bypass the unneeded computation of unnecessary components, thus 341 saving computational time. 342

343 **Parametric Programming** In the Theorem 1, the truncated intervals \mathcal{Z} can be regarded as the 344 intersections of the polytopes $\{P_k^A\}_{k:A_k=A_x}$ with the line $X = \mathcal{Q}_x + \Sigma \eta (\eta^\top \Sigma \eta)^{-1} Z$. This implies 345 that determining the truncated intervals $\mathcal Z$ is accomplished by examining this specific line rather 346 than the entire space. Alogorithm 1 outlines the procedure to identify Z. The algorithm starts at 347 z_{\min} and search for the truncated intervals along the line until z_{\max}^2 . For each step, given z, the algorithm computes the lower bound L_k^A and upper bound U_k^A of the interval to which z belongs 348 349 to, as well as corresponding anomaly region $A_k = A_{\mathbf{X}(z)}$. The $L_k^{\mathcal{A}}$ and $U_k^{\mathcal{A}}$ are computed by the 350 technique described in the next subsection. This procedure is commonly referred to as parametric 351 programming known as parametric programming, which is a method to solve the optimization 352 problem for parameters such as the lasso regularization path (Efron et al., 2004; Hastie et al., 2004; Karasuyama et al., 2012).

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> **Auto-Conditioning** In line 4 of Algorithm 1, we utilize a technique referred to as *auto-conditioning*. 355 Similar to *auto-differentiation*, this method leverages the fact that the entire computations of $L_k^{\mathcal{A}}$ 356 and U_k^A executes a sequence of piece-wise linear operations. By applying the recursive rule re-357 peatedly to these operations, $L_k^{\mathcal{A}}$ and $U_k^{\mathcal{A}}$ can be automatically computed. The details are deferred 358 to Appendix A.3. This implies that by implementing the computational techniques for known 359 piecewise-linear/assignment functions, we can automatically compute the truncation intervals and 360 the anomaly region. This adaptability proves particularly advantageous when dealing with complex 361 systems like Deep Neural Networks (DNNs), where frequent and detailed structural adjustments are 362 often required. We note that the auto-conditioning technique is originally proposed in Miwa et al. (2023). However, the authors concentrate on a specific application of the saliency region, and no 364 existing studies recognize its crucial application in VAE literature. In this paper, we prove that a VAE 365 can be represented as a piecewise-assignment function, thus highlighting the crucial application of 366 auto-conditioning in efficiently conducting the proposed VAE-AD Test.

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6 EXPERIMENT

We demonstrate the performance of the proposed method. More details and results can be found in the Appendix A.5.

Experimental Setup. We compared the proposed method (VAE-AD Test) with OC (simple extension of SI literature to our setting), Bonferroni correction (Bonf) and naive method. More details can be found in Appendix A.5. We considered two covariance matrix structures:

²We set the $z_{\min} = -|T(x)| - 10\sigma$ and $z_{\max} = |T(x)| + 10\sigma$, where σ is the standard deviation of test statistic. This is justified by the fact that the probability in the tails of the normal distribution can be considered negligible.

378 Algorithm 1 Parametric Programming-Based SI 379 **Require:** x, z_{\min}, z_{\max} 380 1: Obtaine $A_{\boldsymbol{x}}$ and compute $\boldsymbol{\eta}$. 381 2: $z \leftarrow z_{\min}$ and $\mathcal{Z} \leftarrow \emptyset$ 382 3: while $z \leq z_{\max}$ do 4: Compute $L_k^{\mathcal{A}}, U_k^{\mathcal{A}}$, and A_k respect to z by *auto-conditioning* (see Appendix A.3). 384 5: if $A_k = A_x$ then $\mathcal{Z} \leftarrow \mathcal{Z} \cup [L_k^{\mathcal{A}}, U_k^{\mathcal{A}}]$ 6: 386 7: end if 387 $z \leftarrow U_k^{\mathcal{A}} + \delta$, where δ is a small positive number. 8: 388 9: end while 10: $p_{\text{selective}} \leftarrow \text{equation 11}$ with Threorem 1 389 output $p_{\text{selective}}$ and $A_{\boldsymbol{x}}$ 390 391 392 393 394 396 397



Figure 2: Type I errors (false positive detection rates) and powers (true positive detection rates) of the proposed VAE-AD Test and three baselines, Naive, OC and Bonf in Indepence and Correlation setting. Naive test, which does not consider the fact that abnormal regions are selected in a data-driven manner, fails to control the Type I error, failing to meet the requirements of a statistical test. On the other hand, the proposed method, VAE-AD Test, and two other baselines, OC and Bonf, all successfully control the Type I error at 0.05 in all settings. The power of the proposed VAE-AD Test is significantly larger than two baselines, OC and Bonf in all problem settings.

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• $\Sigma = I_n$ (Independence)

412 413 • $\Sigma = AR(1) \otimes AR(1)$ (Correlation) where AR(1) is the first-order autoregressive matrix 414 $\{AR(1)\}_{ij} \in \mathbb{R}^{\sqrt{n} \times \sqrt{n}} = 0.25^{|i-j|}$ and \otimes is kronecker dot.

To examine the type I error rate, we generated 1000 null images $X = (X_1, \ldots, X_n)$, where s = 0and $\epsilon \sim N(\mathbf{0}, \Sigma)$, for each $n \in \{64, 256, 1024, 4096\}$. To examine the power, we set n = 256 and generated 1000 images in which $\epsilon \sim N(\mathbf{0}, \Sigma)$, the signals $s_i = \Delta$ for any $i \notin S$ where S is the "true" anomaly region whose location is randomly determined, and $s_i = 0$ for any $i \notin S$. We set $\Delta \in \{1, 2, 3, 4\}$. In all experiments, we set the threshold $\lambda = 1.2$ for the anomaly detection, and the significance level $\alpha = 0.05$. We also apply mean filtering to the reconstruction error to enhance the anomaly detection performance.

422 Numerical results. The results of type I error rate and power are shown in Fig. 2. The VAE-AD 423 Test, OC, and Bonf successfully controlled the type I error rate in the both cases of independence and 424 correlation, whereas the naive method could not. Since the naive method failed to control the type I 425 error, we no longer considered its power. The power of the VAE-AD Test was the highest among the 426 methods that controlled the type I error. The Bonferroni method has the lowest power because it is 427 conservative due to considering the huge number of all possible hypotheses. OC also has low power 428 because it considers extra conditioning, which causes the loss of power.

Real data experiments. We examined the brain image dataset extracted from Buda et al. (2019),
which includes 939 and 941 images with and without tumors, respectively. The results of statistical
testing for images without tumor and with tumor are presented in Figs. 3 and 4. The naive *p*-value is
small even in cases where no tumor region exists in the image. This indicates that the naive *p*-value



Figure 3: Anomaly detection for images without tumor. The naive *p*-values are 0.000 in all settings, incorrectly detecting abnormalities. However, the selective *p*-values based on the proposed VAE-AD Test are all large enough, correctly identifying the absence of abnormalities.



Figure 4: Anormaly detection for images with tumor. In all settings, both the naive *p*-values and the selective *p*-values are low, correctly identifying the abnormalities (although naive *p*-values are invalid statistical tests because it fails to control type I errors).

cannot be used to quantify the reliability of the result of anomaly detection using VAE. With the proposed selective *p*-values, we successfully identified false and true positive detections.

7 CONCLUSIONS, LIMITAIONS AND FUTURE WORKS

We introduced a novel statistical testing framework for AD task using deep learning model. We developed a valid statistical test for VAE-based AD using CSI. We believe that this study stands as a significant step toward reliability of deep learning model-based decision making. There are several constraints on the class of problems where CSI can be applied, so new challenges arise when applying VAEs to other types of neural networks. Additionally, we selected simple options for defining the anomalous region and the test statistic, but it is unknown whether the same framework can be applied to more complex options. Furthermore, as the size of the VAE network increases, the computational cost of calculating the selective *p*-value also increases, necessitating the development of cost reduction methodologies such as parallelization.

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A APPENDIX

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611 A.1 THE DETAILS OF VAE

613 We used the architecture of the VAE as shown in Figure 5 and set m = 10 as a dimensionality of the 614 latent space. We used ReLU as an activation function for the encoder and decoder. We generated 615 1000 images from $N(\mathbf{0}, I_n)$ as normal images and trained the VAE with these images, and used 616 Adam Kingma & Ba (2017) as an optimizer.



Figure 5: Architecture of the VAE.

A.2 PROOF OF THEOREM 1

Proof. The theorem is based on the Lemma 3.1 in Chen & Bien (2020). By the definition of the piecewise-assignment function, the conditional part, $\{M_X = M_x\}$ can be characterized as the union of polytopes,

$$\{M_{\boldsymbol{X}} = M_{\boldsymbol{x}}\} = \bigcup_{k:M_k = M_{\boldsymbol{x}}} P_k^{\mathcal{M}}$$

By substituting $X(Z) = Q_x + \Sigma \eta (\eta^\top \Sigma \eta)^{-1} Z$ into the polytopes P_k^M , we obtain the truncated intervals Z in the lemma. For the set k such that $M_k = M_x$, we have $Q_X \perp Z$ by orthogonality of Q_X and η and by the properties of the normal distribution. Hence, we obtain

$$Z \mid \{M_{\mathbf{X}} = M_{\mathbf{x}}, \mathcal{Q}_{\mathbf{X}} = \mathcal{Q}_{\mathbf{x}}\} \stackrel{d}{=} Z \mid \{Z \in \mathcal{Z}, \mathcal{Q}_{\mathbf{X}} = \mathcal{Q}_{\mathbf{x}}\}$$
$$\stackrel{d}{=} Z \mid \{Z \in \mathcal{Z}\} (:: \mathcal{Q}_{\mathbf{X}} \perp Z)$$

645 There is no randomness in \mathcal{Z} ,

$$Z \mid \{M_{\boldsymbol{X}} = M_{\boldsymbol{x}}, \mathcal{Q}_{\boldsymbol{X}} = \mathcal{Q}_{\boldsymbol{x}}\} \sim TN(\boldsymbol{\eta}^{\top}\boldsymbol{\mu}, \boldsymbol{\eta}^{\top}\boldsymbol{\Sigma}\boldsymbol{\eta}; \mathcal{Z}).$$

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648 A.3 THE DETAILS OF AUTO-CONDITIONING

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This section demonstrates the auto-conditioning algorithm, utilized to compute the truncated intervals $[L_k^A, U_k^A]$ and the corresponding anomaly region A_k respect to the z in Algorithm 1. The algorithm is introduced for the piecewise-assignment function, which is composed of piecewise-linear functions and a piecewise-assignment function.

654 It is conceptualized as a directed acyclic graph (DAG) that delineates the processing of input data, 655 similar to a computational graph in auto-differentiation. In this graph, the nodes symbolize the 656 piecewise-linear and piecewise-assignment functions, each with an input and output edge to represent 657 the function compositions. It should be noted that the node such as μ_{ϕ} and μ_{θ} , may replace the other DAG express the piecewise-linear/assignment function of the node since it can be represented as the 658 composition and concatenation of array of simpler piecewise-linear/assignment functions. The level 659 of simplicity for a function of a node can be determined based on what is most convenient for the 660 implementation. A special node, representing the concatenation of two piecewise-linear functions, 661 features two input edges and one output edge. Figure 6 shows the directed acyclic graph of the 662 anomaly detection using VAE in Eq. equation 4. 663



Figure 6: The directed acyclic graph of the anomaly detection using VAE $\mathcal{A} : \mathbb{R}^n \to 2^{|n|}$ defined in Eq. equation 4. Circles represent the piecewise-linear functions and the piecewise-assignment function. The rectangle represents the concatenation of piecewise-linear functions. The edges represent the composition of piecewise-linear functions.

A.3.1 UPDATE RULES FOR THE NODES OF THE PIECEWISE-ASSIGNMENT FUNCTIONS

The computation of the interval $[L_k^A, L_k^A]$ is defined in a recursive way. The output of the node $f: \mathbb{R}^l \to \mathbb{R}^m$ in the DAG are denoted as $a_f, b_f \in \mathbb{R}^m$ and $L_f, U_f \in \mathbb{R}$.

Update rule for the initial node. At first, the output of the initial node X(z) of the directional graph denoted as f_0 for notational convention, are defined as $a_{f_0} = Q_x$, $b_{f_0} = \Sigma \eta (\eta^\top \Sigma \eta)^{-1}$, $L_{f_0} = -\infty$, and $U_{f_0} = \infty$. It should be noted here that $X(z) = a_{f_0} + b_{f_0} z$ is the line appeared in the proof of Theorem 1 in Section A.2.

688 **Update rule for the node of the piecewise-linear functions.** Let us consider the output for the 689 node q whose input is the output of the node f in the DAG. The inputs of the q's node (i.e. output of 690 node f) are denoted as a_f, b_f, L_f and U_f . a_f is the summed point vector added in the piecewise-691 linear functions until reaching f, b_f is the direction vector corresponding to z, multiplied in the 692 piecewise-linear functions until reaching to f. Then, the output of the piecewise-linear function f is 693 represented as $a_f + b_f z$. L_f and U_f are the lower and upper bounds of the interval obtained at the 694 piecewise-linear function f. The output of the node g is defined as follows: 1) Check the index jsuch that the output of f within the polytope of: $P_i^g \ni a_f + b_f z$. 2) Compute the point vector a_q and the direction vector \boldsymbol{b}_q of the piecewise-linear function g with the index j, 696

$$\boldsymbol{a}_g = \boldsymbol{\Psi}_j^g \boldsymbol{a}_f + \boldsymbol{\psi}_j^g, \ \boldsymbol{b}_g = \boldsymbol{\Psi}_j^g \boldsymbol{b}_f. \tag{13}$$

699 3) Compute the lower and upper bounds of the interval L_g and U_g with the index j,

$$L = \max_{k: (\beta_j^g)_k > 0} \frac{(\boldsymbol{\alpha}_j^g)_k}{(\beta_j^g)_k}, \ U = \min_{k: (\beta_j^g)_k < 0} \frac{(\boldsymbol{\alpha}_j^g)_k}{(\beta_j^g)_k}$$

where $\alpha_j^g = \delta_j^g - \Delta_j^g a_f$ and $\beta_j^f = \Delta_j^g b_f$. 4) Take the intersection of the interval $[L_f, U_f] \cap [L, U]$ as the interval $[L_g, U_g]$ of the piecewise-assignment function g as

$$L_q = \max(L_f, L), \ U_q = \min(U_f, U).$$

This update rule is obtained from the Lemma 2 in Miwa et al. (2023).

Update rule for the nodes of concatenation of two piecewise-linear functions. Let us consider the concatenation node of two piecewise-linear functions f and g denoted as concat. Let the inputs of the node be a_f, b_f, L_f and U_f from the node f and a_g, b_g, L_g and U_g from the node g. The output of the concatenation node, $a_{concat}, b_{concat}, L_{concat}$ and U_{concat} are defined as follows: 1) Concatenate the vector outputs of nodes f and g

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 $oldsymbol{a}_{ ext{concat}} = egin{bmatrix} oldsymbol{a}_f \ oldsymbol{a}_g \end{bmatrix}, \ oldsymbol{b}_{ ext{concat}} = egin{bmatrix} oldsymbol{b}_f \ oldsymbol{b}_g \end{bmatrix}.$

2) Take intersection of the interval $[L_f, U_f] \cap [L_g, U_g]$ as

 $L_{\text{concat}} = \max(L_f, L_g), \ U_{\text{concat}} = \min(U_f, U_g).$

T19 Update rule for the final node. At the final node Θ which is the piecewise-assignment function, it takes the same input as the node of piecewise-linear functions and outputs are the same except for the a_{Θ} and b_{Θ} . 1) It computes the index j such that the input falls into the polypotopes of P_j^{Θ} . 2) Then, the anomaly region A_j is obtained instead of Eq. equation 13 in the update rule for the node of piecewise-linear functions. 3) The computation of lower bounds L_{Θ} and the upper bounds U_{Θ} are the same as the update rule for the node of piecewise-linear functions. The output of the final node are the anomaly region A_j , the lower bounds L_{Θ} and the upper bounds U_{Θ} .

Then, apply the above update rule to the directional graph of the piecewise-assignment function from the initial node f_0 to the final node Θ . Consequently, the auto-conditioning algorithm computes the lower and upper bounds of the interval as the outputs of final node $L_k^A = L_{\Theta}, L_k^A = L_{\Theta}$ and $A_k = A_j$.

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731 A.4 IMPLEMENTATION

We implemented the auto-conditioning algorithm described above in Python using the tensorflow library. The codes construct the DAG of the piecewise-assignment function automatically from the trained Keras/tensorflow model. Then, we do not need further implementation to conduct CSI for each specific DNN model. This indicates that even if we change the architecture or adjust the hyper-parameters and retrain the DNN models, we can conduct the CSI without additional implementation.

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A.5 EXPERIMENTAL DETAILS

Methods for comparison. We compared our proposed method with the following methods:

• VAE-AD Test: our proposed method.

• OC: our proposed method conditioning on the only one polytope to which the observed image belongs $x \in P_k^A$. This method is computationally efficient; however, its power is low due to over-conditioning.

• Bonf: the number of all possible hypotheses are considered to account for the selection bias. The *p*-value is computed by $p_{\text{bonf}} = \min(1, p_{naive} \times 2^n)$

• Naive: the conventional method is used to compute the *p*-value.

Experiment for robustness. We evaluate the robustness of our proposed methodology in terms of Type I error control, specifically under conditions where the noise distribution deviates from the Gaussian assumption. We investigate this robustness by applying our method across a range of non-Gaussian noise distributions, including:

• Skew normal distribution (*skewnorm*)

• Exponential normal distribution (exponorm)

• Generalized normal distribution with steep tails (gennormsteep)

- Generalized normal distribution with flat tails (gennomflat)
- Student's t distribution (t)

We commence our analysis by identifying noise distributions from the aforementioned list that have a 1-Wasserstein distance of $\{0.01, 0.02, 0.03, 0.04\}$ relative to the standard normal distribution N(0, 1). Subsequently, we standardize these noise distributions to ensure a mean of 0 and a variance of 1. Setting the sample size to n = 256, we generate 1000 samples from the selected distributions and apply hypothesis testing to each sample to obtain the Type I error rate. This process is conducted at significance levels $\alpha = \{0.05, 0.10\}$. The results are shown in Fig. 7. Our method still maintains good performance in type I error rate control.



Figure 7: Robustness of type I error control.

More results on brain image dataset. Additional results are shown in Figs. 8 and 9.

Computational resources used in the experiments. All numerical experiments were conducted on a computer with a 56-core 2.00GHz CPU, eight RTX-A6000 GPUs, and 1024GB of memory.



