
Solving dynamic portfolio selection problems via score-based diffusion models

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Abstract

In this work, we tackle the dynamic mean-variance portfolio selection problem in a *model-free* manner, based on (generative) diffusion models. We propose using data sampled from the real model \mathbb{P} (which is unknown) with limited size to train a generative model \mathbb{Q} (from which we can easily and adequately sample). With adaptive training and sampling methods that are tailor-made for time-series data, we obtain quantification bounds between \mathbb{P} and \mathbb{Q} in terms of the adapted Wasserstein metric \mathcal{AW}_2 . We then propose a policy gradient algorithm based on the generative environment, in which our innovative adapted sampling method provides approximate scenario generators. We illustrate the performance of our algorithm on real data. The algorithm based on the generative environment produces portfolios that beat several important baselines, including the Markowitz portfolio, the equal weight (naive) portfolio, and S&P 500.

1 Introduction

Portfolio selection, a crucial and fundamental task in financial engineering, has been a subject of both theoretical and practical interest since the classical work Markowitz [1952]. With the ability to account for intertemporal hedging and rebalancing in a non-stationary market, formulating portfolio selection as *dynamic* multi-period problems has become increasingly appealing; see Mossin [1968], Merton [1971] for early works. Due to incompleteness of the information, estimation errors, or simply the infeasibility of fully modeling market uncertainty, we lack access to the real model, as opposed to model-based approaches. Instead, we are restricted to a limited set of financial data samples. Consequently, investigating portfolio selection problems in a *model-free* manner has been a long-standing topic in financial engineering; see discussions of ‘universal portfolio’ Cover [1991], stochastic portfolio theory Karatzas and Ruf [2017], Cuchiero et al. [2019], and the application of machine learning Ban et al. [2018], Gu et al. [2020]. However, dynamic portfolio selection problems are rather involved, even with model-based approaches, due to their inherent intertemporal structure Campbell and Viceira [1999], high dimensionality Brandt et al. [2005], and time inconsistency Basak and Chabakauri [2010], Björk et al. [2014]. Therefore, model-free dynamic portfolio selection problems are worth exploring.

To address this challenge, we propose to leverage score-based diffusion models. These models can generate additional samples that exhibit distributional properties similar to the original data, but theoretical foundations for generating time-series data have not been established yet. We close this gap by proposing a adaptive sampling scheme (Algorithm 1), which outputs an alternative model \mathbb{Q} . The difference between \mathbb{Q} and the original model \mathbb{P} is assessed with adapted Wasserstein metric (Theorem 2.1). We then utilize the model \mathbb{Q} as an oracle in the subsequent policy gradient algorithm: we sample a large number of paths $\bar{s}^{1:T} \sim \mathbb{Q}$ and train a reinforcement learning (RL) agent using these simulated paths. The approximated model \mathbb{Q} consists of a score network s_θ , which is the core of the score-matching technique used in generative models, and a Recurrent Neural Network (RNN) encoder R_θ , which encodes the current state of the market. These two networks are trained simultaneously using historical price data. The reinforcement learning agent is trained on the generated data from \mathbb{Q} , which are produced by evaluating s_θ and R_θ in an adaptive manner. Finally, the agent, fed with the encoded (and updated) market state, outputs portfolio selection actions whose performance is evaluated through real-world experiments.

2 Score-based diffusion model for time-series data

For simplicity, we take the OU process as the forward process of our diffusion model:

$$dX_\tau = -X_\tau d\tau + \sqrt{2}dB_\tau. \quad (1)$$

The corresponding time-reversed SDE is

$$d\bar{X}_\tau = [\bar{X}_\tau + 2\nabla \log p(\mathcal{T} - \tau, \bar{X}_\tau)]d\tau + \sqrt{2}d\bar{B}_\tau,$$

where p is the density function of the forward process (1) with certain initial conditions. For the purpose of conditional score-matching and sampling, we use $p_{t+1}(\tau, \cdot | x^{1:t})$ to denote the probability density function of the forward process, with random initial condition $\mathbb{P}_{x^{1:t}}$. Here, $x^{1:t}$ is the conditional input and can be used in practice as observations.

Suppose the real model of time-series data is \mathbb{P} , and for every t we have a pre-trained diffusion model s_θ^t with a score matching error $\varepsilon_{\text{score}}$ to the real (conditional) score function $\nabla_x p_t$ (details of the conditional score matching will be provided in Appendix B). We note that, fixing $x^{1:t}$, $s_\theta^{t+1}(\tau, x^{1:t}, \cdot)$, serves as an approximation of the *conditional score function* $\nabla_x \log p(\tau, \cdot | x^{1:t})$. This motives us to sample from the output distribution adaptively; see Algorithm 1.

Algorithm 1 Adaptive sampling

Inputs: pre-trained approximated score functions s_θ^t , $t = 1, 2, \dots, T$; backward SDE simulators $\{Y_s\}_{0 \leq s \leq \mathcal{T}}$.

Initialization: samples from the noise: $\{(Z_{(n)}^1, Z_{(n)}^2, \dots, Z_{(n)}^T)\}_{n=1}^N \sim \mathcal{N}(\mathbf{0}_{\mathbb{R}^{dT}}, I_{dT \times dT})$.

1: **for** $n = 1$ **to** N **do**

2: With initial condition $\bar{Y}_0 = Z_{(n)}^1$, run the backward SDE with score matching function s_θ^1 , i.e.,

$$d\bar{Y}_\tau = [\bar{Y}_\tau + 2s_\theta^1(\mathcal{T} - \tau, \bar{Y}_\tau)]d\tau + \sqrt{2}\bar{B}_\tau.$$

3: $y_{(n)}^1 \leftarrow \bar{Y}_\mathcal{T}$.

4: **for** $t = 1$ **to** $T - 1$ **do**

5: With initial condition $\bar{Y}_0 = Z_{(n)}^{t+1}$, run the backward SDE with score matching function $(\tau, x) \mapsto s_\theta^{t+1}(\tau, y_{(n)}^{1:t}, x)$, i.e.,

$$d\bar{Y}_\tau = [\bar{Y}_\tau + 2s_\theta^{t+1}(\mathcal{T} - \tau, y_{(n)}^{1:t}, \bar{Y}_\tau)]d\tau + \sqrt{2}\bar{B}_\tau.$$

6: $y_{(n)}^{t+1} \leftarrow \bar{Y}_\mathcal{T}$.

7: **end for**

8: **end for**

Output: $\{(y_{(n)}^1, y_{(n)}^2, \dots, y_{(n)}^T)\}_{n=1}^N$.

We now present our main theoretical result: an quantification error bound between the real model \mathbb{P} and the output measure $\mathbb{Q}^\mathcal{T}$ of Algorithm 1.

Theorem 2.1. For any $\varepsilon_{\text{score}} > 0$, there exists a family of score matching functions $\{s_\theta^t\}_{t=1}^T$ with score matching error $\varepsilon_{\text{score}}$, such that Algorithm 1 produces an approximated model \mathbb{Q}^T satisfying

$$AW_2^2(\mathbb{P}, \mathbb{Q}^T) \leq C \left(\mathcal{T}^{\frac{5\mathcal{T}}{2}} \varepsilon_{\text{score}}^{1/2\mathcal{T}-1} + \mathcal{T}^{\frac{5(\mathcal{T}-1)}{2} + \frac{1}{2\mathcal{T}-2}} e^{-c\mathcal{T}/2\mathcal{T}-1} \varepsilon_{\text{score}}^{-1/2\mathcal{T}-1} \right). \quad (2)$$

Remark 1. Unlike usual Wasserstein bound in static data setting, the error is unbounded when we let $\varepsilon_{\text{score}} \rightarrow 0$ for a fixed \mathcal{T} . However, by taking $\varepsilon_{\text{score}} \rightarrow 0$ and $\mathcal{T} \rightarrow \infty$ in an appropriate scale (e.g., $\varepsilon_{\text{score}} = \mathcal{T}^{-52\mathcal{T}-3+1} e^{-c\mathcal{T}/2}$), we still obtain small error bounds. This would indeed be the case when we implement finite sample analysis, where $\varepsilon_{\text{score}}$ and \mathcal{T} are in different scales of sample size n . Such a difference arises because we deviate from usual strongly log concavity assumption of data distribution. This assumption seems rather restrictive when imposed on all *conditional* distributions. See technical appendices for details.

3 Application to portfolio selection

In this section, we apply score-based diffusion model, redesigned for time-series data in this work, to dynamic portfolio selection problem. More specifically, suppose the price data $S^{1:T} \sim \mathbb{P}^1$, we want to solve

$$v^*(\mathbb{P}) := \sup_{\vartheta} \left\{ \mathbb{E}_{\mathbb{P}}[(\vartheta \cdot S)_T] - \frac{\gamma}{2} \text{Var}_{\mathbb{P}}[(\vartheta \cdot S)_T] \right\}, \quad (3)$$

where $(\vartheta \cdot S)_T := \sum_{l=1}^{T-1} \vartheta_l^t (S^{l+1} - S^l)$. The idea is to solve (3) under \mathbb{Q} (i.e., $v^*(\mathbb{Q})$) by a policy gradient algorithm, and use the output portfolio as an approximated solution to (3). This is feasible because by results in Section 2 we can sample from \mathbb{Q} easily and the conditional sampling is also allowed.

To implement Algorithm 1, we use two neural networks to represent \mathbb{Q} : one score network s_θ and one RNN encoder R_θ . The RNN encoder is used to summarize the information from history by the recursive relation $h^{t+1} = R_\theta(s^t, h^{t-1})$, hence the conditional variables $s^{1:t}$ with varying length can be replaced by h^t on a single space. s_θ and R_θ are trained jointly by minimizing the following score-matching loss function:

$$\frac{1}{TM} \sum_{t,m} \left| \sqrt{1 - e^{2\tau(m)}} s_\theta \left(\tau(m), h_{(m)}^t, s_{(m)}^t e^{-\tau(m)} + \sqrt{1 - e^{-2\tau(m)}} \mathbf{z}_{(m)} \right) + \mathbf{z}_{(m)} \right|^2. \quad (4)$$

Here, $\{s_{(m)}^{1:T}\}_{m=1}^M$ is a batch of data, and $\{h_{(m)}^{1:T}\}_{m=1}^M$ are the corresponding hidden variables obtained by feeding the data to the RNN encoder R_θ ; $\{\tau(m)\}_{k=1}^M$ are randomly sampled from $[0, \mathcal{T}]$; $\{\mathbf{z}_{(m)}\}_{m=1}^M$ are sampled from pure noise $\mathcal{N}(0, I)$. Although only s_θ explicitly appears in (4), R_θ is implicitly trained by backward propagation. In the sampling phase, we follow Algorithm 1, while replacing $y_{(n)}^{1:t}$ obtained from previous steps by $h_{(n)}^t = R_\theta(s_{(n)}^t, h_{(n)}^{t-1})$. We emphasize that the proposed sampling algorithm support *conditional sampling*: we can collect some contextual data $x^{-1:-T}$ and use it to obtain h^0 . Inserting h^0 into s_θ enables sampling $s^{1:T}$ from some conditional distribution instead of a fixed distribution \mathbb{P} . This will be crucial for applications to portfolio selection. An implemented version of Algorithm 1, along with the training algorithm, is provided in Appendix E as pseudocode; see Algorithms 2 and 3.

As the next step, we train a reinforcement learning agent with a policy gradient algorithm with the *generative environment* \mathbb{Q} . The proposed algorithm is a revised version of the deterministic policy gradient algorithm (DDPG, c.f. Lillicrap et al. [2016]). We redesign part of the algorithm to accommodate the unique characteristics of portfolio selection problems. These adjustments include separating the scenario pool and replay buffer, which facilitates the convenient integration of the generative model. See Algorithm 4 in Appendix E for the pseudo code.

4 Experiments

We implement the proposed workflow to a set of real world data. Our empirical tests employ the value-weighted 10 Industry Portfolios from the Kenneth R. French Data Library, a monthly panel that aggregates all common stocks on the CRSP tape into ten broad industry groups—Non-Durables (‘NoDur’),

¹We change the notation to S as is frequently used in finance

Table 1: Performance of different portfolio strategies with $\gamma = 0.5$

	Return	Volatility	Sharpe	Sortino	Max Drawdown	Calmer
S&P 500	10.98%	0.1460	0.7522	0.6955	-0.2477	0.4433
EW	13.26%	0.1502	0.8831	0.8310	-0.2293	0.5783
HistMarkowitz	6.18%	0.2043	0.3025	0.2973	-0.4090	0.1511
GenMarkowitz	12.02%	0.2418	0.4970	0.5121	-0.5563	0.2160
GenTD3	13.57%	0.1440	0.9428	0.9774	-0.2199	0.6405

Durables ('Durbl'), Manufacturing ('Manuf'), Energy ('Enrgy'), HiTech ('HiTec'), Telecommunications ('Telcm'), Utilities ('Utils'), Shops ('Shops'), Health ('Hlth'), and Others ('Other'). The dataset that we use can be found and downloaded from https://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html. We provide it in the supplementary files along with the code.

For training, we use a *same-resolution UNet* which is inspired by Ronneberger et al. [2015] as the score network s_θ and an LSTM Hochreiter and Schmidhuber [1997] as the RNN encoder R_θ like Yan et al. [2021]. To improve the performance, we employ variance-preserving stochastic differential equation (VPSDE) with a time varying noisy scheme $\{\beta(\tau)\}_{\tau \in [0, T]}$, which generalizes the OU process in theoretical part. Details of training, including choices of hyperparameters, can be found in Appendix D. For the RL gradient, we use the *twin delayed* version of Algorithm 4 (Twin Delayed DDPG, **TD3**; c.f. Fujimoto et al. [2018]) to improve performance and stability. The training (along with validation) dataset consists of indices data from July 1926 through March 2009.

For testing, we set the time horizon $T = 12$ (i.e., one year) for TD3 agent, and use an extended test time horizon $T_{\text{test}} = 15T$ (i.e., 15 years) to show the robustness. That is to say, we use indices data spanning from April 2009 to March 2025 for the purpose of testing. We compare the market benchmark, the S&P 500 index, and four portfolio strategies: Equal Weight (EW), History-based Markowitz (HistMarkowitz), Generative-model-based Markowitz (GenMarkowitz), and Generative-model-based TD3 (GenTD3). The results are presented in Table 1. Here, we set the risk aversion $\gamma = 0.5$. More experiment results, including those with varying risk aversions, are provided in Appendix D.

5 Limitations and future works

As a main theoretical limitations, the Lipschitz continuity on the conditional variable $x^{1:t}$ in (10) seems restrictive. One could relax this assumption with a more refined approximation analysis, impose Lipschitz continuity only on x^{t+1} , which aligns with the usual Lipschitz condition for the score function. Moreover, the error bounds established here are unlikely to be optimal; tightening, or even optimizing, the rate remains an interesting direction. *Practically*, further fine-tuning of the proposed algorithms and large-scale experiments are highly motivated. Furthermore, to enhance the predictive ability we can incorporate more covariant variables into the input of RNN encoder, instead of just using the price (return) data. Quantitative analysis of the RL part of this work (e.g, the regret analysis) is also interesting. We leave these directions for future study.

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A Frequently used notation

In this work, the data is in the form of time-series with length $T \in \mathbb{N}$ and dimension $d \in \mathbb{N}$. We denote by $t \in \{1, 2, \dots, T\}$ the time index for the time-series. For general theoretical results, we use $X^{1:t} := (X^1, X^2, \dots, X^t)$ ($Y^{1:t} := (Y^1, Y^2, \dots, Y^t)$) to denote random variables sampled from data distribution (approximated model, respectively), and $x^{1:t}$ ($y^{1:t}$) the realization. For the diffusion model, we use $\tau \in [0, T]$ to denote the *diffusive time*, i.e., the time flow of the forward and reversed SDE, with $T > 0$ the maximal diffusive time.

For a probability \mathbb{P} on the space of time-series data \mathbb{R}^{dT} , $\mathbb{P}_{1:t}$ is the joint distribution of (X^1, X^2, \dots, X^t) , thus $\mathbb{P}_{1:T} = \mathbb{P}$. Moreover, $\mathbb{P}_{x^{1:t}}$ is the conditional distribution of X^{t+1} , conditioning on $X^{1:t} = x^{1:t}$. By convention, we use \mathbb{P} to denote the original (real) model, hence the distribution of $X^{1:T}$ and \mathbb{Q} the alternative model, hence the distribution of $Y^{1:T}$.

To present the quantification error bound of the score-based diffusion model, we use the adapted Wasserstein metric. More specifically, for two probability measures \mathbb{P} and \mathbb{Q} on \mathbb{R}^{dT} , a coupling π between \mathbb{P} and \mathbb{Q} is said to be *causal in x* if for any $t \in \{1, 2, \dots, T\}$,

$$\pi(Y^t \in dy^t | X^{1:T} = x^{1:T}) = \pi(Y^t \in dy^t | X^{1:t} = x^{1:t}).$$

A coupling π is said to be *bicausal* if it is causal in both x and y (c.f. Backhoff et al. [2017], Bayraktar and Han [2025+]). Define $\Pi_{bc}(\mathbb{P}, \mathbb{Q})$ to be the set of all bicausal couplings between \mathbb{P} and \mathbb{Q} . With the choice of transport cost $|x^{1:T} - y^{1:T}| := (\sum_{t=1}^T |x^t - y^t|^2)^{1/2}$, we define the *adapted Wasserstein metric* as follows:

$$AW_2^2(\mathbb{P}, \mathbb{Q}) = \left(\inf_{\pi \in \Pi_{bc}} \int_{\mathbb{R}^{dT} \times \mathbb{R}^{dT}} |x^{1:T} - y^{1:T}|^2 \pi(dx^{1:T}, dy^{1:T}) \right)^{1/2}.$$

We use K, C, c to denote generic constants which may vary from line to line.

B Details of the score matching

In Section 2, we assume that the approximated score functions $\{s_\theta^t\}_{t=1}^T$ have the score matching error $\varepsilon_{\text{score}}$. More specifically, we assume that they satisfy the following assumption:

Assumption 1. For any $\tau \in (0, T]$ and $t = 1, 2, \dots, T-1$, we have

$$\mathbb{E}_{X \sim p_1(\tau, \cdot)} |s_\theta^1(\tau, X) - \nabla_x \log p_1(\tau, X)|^2 \leq \varepsilon_{\text{score}}^2, \quad (5)$$

$$\mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) - \nabla_x \log p_{t+1}(\tau, X_\tau^{t+1} | X^{1:t}) \right|^2 \leq \varepsilon_{\text{score}}^2 \quad (6)$$

Here, $p_{t+1}(\tau, \cdot | X^{1:t})$ is the probability density function of the forward process, with random initial condition $\mathbb{P}_{x^{1:t}}$.

We remark that (5) is the classical score error bound for the first marginal of \mathbb{P} ; see Song et al. [2021b]. (6) is crucial to our adaptive sampling scheme because it provides a proxy to $\nabla_x \log p_{t+1}(\tau, \cdot | X^{1:t})$, the score function conditioning on the realization $X^{1:t}$. However, (6) is not feasible directly, because we can not sample from the conditional probability, but rather, only have access to the joint distribution (data distribution). This issue can be resolved by the following *denoising score-matching techniques*, in which the error can be computed using only joint distribution. While the idea is borrowed from classical denoising score matching (e.g., Hyvärinen [2005]), we need to consider a conditional version and the proof is more complicated.

Proposition B.1. For any $t = 1, 2, \dots, T - 1$, the following two score-matching problems are equivalent:

$$\min_{\theta} \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_{\tau}^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_{\theta}^{t+1}(\tau, X^{1:t}, X_{\tau}^{t+1}) - \nabla_x \log p_{t+1}(\tau, X_{\tau}^{t+1} | X^{1:t}) \right|^2 \quad (7)$$

$$\min_{\theta} \mathbb{E}_{X^{1:t+1} \sim \mathbb{P}_{1:t+1}} \mathbb{E}_{X_{\tau}^{t+1} \sim \phi(\tau, \cdot | X^{t+1})} \left| s_{\theta}^{t+1}(\tau, X^{1:t}, X_{\tau}^{t+1}) - \nabla_x \log \phi(\tau, X_{\tau}^{t+1} | X^{t+1}) \right|^2. \quad (8)$$

Here, $\phi(\tau, \cdot | x_0)$ is the probability density function of the forward process, with initial condition $X_0 = x_0$. In particular,

$$\phi(\tau, x | x_0) = (2\pi(1 - e^{-2\tau}))^{-d/2} e^{-\frac{|x - x_0 e^{-\tau}|^2}{2(1 - e^{-2\tau})}}.$$

Because the explicit form of $\phi(\tau, x | x_0)$, $X_{\tau}^{t+1} \sim \phi(\tau, \cdot | X^{t+1})$ is equivalent to $X_{\tau}^{t+1} = X^{t+1} e^{-\tau} + \sqrt{1 - e^{-2\tau}} \mathbf{z}$, where $\mathbf{z} \sim \mathcal{N}(0, I)$. Therefore, (8) is further equivalent to

$$\min_{\theta} \mathbb{E}_{X^{1:t+1} \sim \mathbb{P}_{1:t+1}} \mathbb{E}_{\mathbf{z} \sim \mathcal{N}(0, I)} \left| \sqrt{1 - e^{-2\tau}} s_{\theta}^{t+1}(\tau, X^{1:t}, X^{t+1} e^{-\tau} + \sqrt{1 - e^{-2\tau}} \mathbf{z}) + \mathbf{z} \right|^2. \quad (9)$$

In the implementation, s_{θ} is trained so as to minimize the empirical version of (9), with an additional integration in τ to account for different diffusive time; see (4).

C Assumptions of Theorem 2.1

To establish Theorem 2.1, we need two technical assumptions, one on data distribution \mathbb{P} and another on approximating score function s_{θ} . We present such assumptions in this appendix, and several remarks are in order.

Assumption 2 (Assumptions on the data distribution).

1. There exists a constant $L > 0$ such that for any $t \in \{1, 2, \dots, T\}$, $\tau \in [0, \mathcal{T}]$ and $x^{1:t}, y^{1:t} \in \mathbb{R}^{dt}$,

$$\begin{aligned} |\nabla_x \log p_1(\tau, x) - \nabla_x \log p_1(\tau, y)| &\leq L|x - y|, \\ |\nabla_x \log p_{t+1}(\tau, x | x^{1:t}) - \nabla_x \log p_{t+1}(\tau, y | y^{1:t})| &\leq L(|x^{1:t} - y^{1:t}| + |x - y|). \end{aligned} \quad (10)$$

2. For some $c > 0$, $\mathbb{E}_{\mathbb{P}_1} e^{c|X_1|} < \infty$, and for $t = \{1, 2, \dots, T - 1\}$,

$$\sup_{x^{1:t} \in \mathbb{R}^{dt}} \mathbb{E}_{\mathbb{P}_{x^{1:t}}} e^{c|X^{t+1}|} < \infty.$$

Assumption 3 (Assumptions on the approximating network).

1. There exists a constant $L > 0$ such that for any $t \in \{1, 2, \dots, T\}$, $\tau \in [0, \mathcal{T}]$ and $x^{1:t}, y^{1:t} \in \mathbb{R}^{dt}$,

$$\begin{aligned} s_{\theta}^1(\tau, x) - s_{\theta}^1(\tau, y) &\leq L|x - y|, \\ s_{\theta}^{t+1}(\tau, x^{1:t}, x) - s_{\theta}^{t+1}(\tau, y^{1:t}, y) &\leq L(|x^{1:t} - y^{1:t}| + |x - y|). \end{aligned}$$

2. There exist constants $M_{\text{disp}}, \delta > 0$ such that for any for any $t \in \{1, 2, \dots, T\}$, $\tau \in [0, \mathcal{T}]$, $x^{1:t} \in \mathbb{R}^{dt}$ and $x \in \mathbb{R}^d$,

$$\begin{aligned} 2x \cdot s_{\theta}^1(\tau, x) &\leq -(1 + \delta)|x|^2 + M_{\text{disp}}, \\ 2x \cdot s_{\theta}^{t+1}(\tau, x^{1:t}, x) &\leq -(1 + \delta)|x|^2 + M_{\text{disp}}. \end{aligned}$$

Remark 2. Typically, it is advisable to refrain from making assumptions about the approximation network, as such assumptions may potentially compromise its approximating power. However, in this paper, we demonstrate that by adhering to Assumption 2, which are all assumptions pertaining to the data distribution, we can select the approximating network s_{θ} such that it satisfies Assumption 3 (with appropriately chosen M_{disp}), while maintaining the approximation errors (5) and (6). See Proposition C.1 below.

Proposition C.1. Suppose Assumption 2 holds. For any $\varepsilon_{\text{score}} > 0$, there exists a s_{θ} satisfying Assumption 1 as well as Assumption 3 with $M_{\text{disp}} \sim \log(1/\varepsilon_{\text{score}})$.

Remark 3. Even for static data ($T = 1$) our theoretical results are new from the perspective of assumptions needed. Indeed, when $T = 1$, Assumption 2 is reduced to the usual Lipschitz continuity of score functions, and we do not rely on any structural assumptions on \mathbb{P} , such as strongly log concavity. The price to pay is that the noise approximation error term $\mathcal{T}^2 e^{-c\mathcal{T}} \varepsilon_{\text{score}}^{-1}$ explodes when $\varepsilon_{\text{score}} \rightarrow 0$. See (2).

D Details and more results of the experiment

As is mentioned in Section 4, in the experiment we use the value-weighted 10 Industry Portfolios from the Kenneth R. French Data Library, spanning July 1926 through March 2025. The detailed split of the dataset is as follows:

1. The test dataset comprises data collected over the past 16 years, from April 2009 to March 2025. The initial year of the test dataset, spanning from April 2009 to March 2008, serves as the historical window and is utilized as the initial condition input into the diffusion model; see variable h in Algorithms 2-4. Consequently, the evaluation period is from March 2010 to March 2025.
2. The validation dataset consists of data from August 1992 to March 2009, encompassing 500 months prior to the test dataset.
3. The remaining dataset is employed in Algorithm 2 to train the diffusion model s_θ and R_θ .

For training, we use a *same-resolution UNet* which is inspired by Ronneberger et al. [2015] as the score network s_θ and an LSTM Hochreiter and Schmidhuber [1997] as the RNN encoder R_θ . To improve the performance, we employ the refined version of OU process: the variance-preserving stochastic differential equation (VPSDE):

$$dX_\tau = -\frac{1}{2}\beta(\tau)X_\tau d\tau + \sqrt{\beta(\tau)}dB_\tau, \quad \beta(\tau) = \beta_{\min} + (\beta_{\max} - \beta_{\min})\tau,$$

discretised into N uniform timesteps $\{t_n\}_{n=1}^N$. Default hyperparameters are $\beta_{\min} = 0.01$, $\beta_{\max} = 10.0$, and $N = 1000$. Then, s_θ and R_θ are trained according to Algorithm 2. Detailed model architectures and training configurations are available in our github repository. Optimisation uses the AdamW optimiser with learning rate 10^{-3} , $(\beta_1, \beta_2) = (0.9, 0.999)$, and weight decay 10^{-2} . A cosine-annealing scheduler runs for 400 epochs with a 5 epoch warm-up. Gradients are clipped to a maximum norm of 1, and mixed precision training (bfloat16) is enabled via `torch.cuda.amp`. We maintain an exponential moving average (EMA) of model weights with decay 0.999 and evaluate validation metrics on the EMA weights. With these settings, training on a single NVIDIA T-40 (On Google Colab) is completed in about 10 minutes.

For testing, we set the time horizon $T = 12$ (i.e., one year) for TD3 agent, and use an extended test time horizon $T_{\text{test}} = 15T$ (i.e., 15 years) to show the robustness. We compare the market benchmark, the S&P 500 index, and four portfolio strategies with details described below:

1. **Equal Weight (EW):** $a \equiv (1/d, 1/d, \dots, 1/d)$;
2. **History-based Markowitz (HistMarkowitz):** For each $t = 1, 2, \dots, T_{\text{test}}$, action a is determined by solving the Markowitz problem with no short selling, no borrowing and fully investment constraints. The mean and covariance are estimated by the sample mean and sample covariance of the most recent 60 months data (as in DeMiguel et al. [2009]). In other words, this is a monthly re-balancing Markowitz strategy in which we use historical data as estimators;
3. **Generative-model-based Markowitz (GenMarkowitz):** For each $t = 1, 2, \dots, T_{\text{test}}$, we collect the most recent T months (i.e., one year) data as the context window, and use it as the initial feature h^0 in the diffusion model (see Algorithm 3). Then, we sample 500 samples of *one month predictions* for s^t . Then, we use this prediction to solve constrained Markowitz problem. In other words, this is a monthly re-balancing Markowitz strategy based on generative model;
4. **Generative-model-based TD3 (GenTD3):** We set the first year returns from the test dataset as the context window $s^{-T:-1}$, and then adaptively sample 500 paths of one-year prediction $s^{1:T}$ based on Algorithm 3. We use these samples as the scenario pool \mathcal{S} to train the TD3 agent. The trained agent is used repeatedly for the first 7.5 years, outputting the portfolio with *updated feature* h^t for any $t = 1, 2, \dots, T_{\text{test}}$. Then, at the midway of the testing period, we use the most recent one-year data as a new context window and generate another set of 500 paths. The agent is retrained using these paths and then used for the rest 7.5 years². The training of the TD3 policy network costs about 30min (including retraining at 7.5 years).

²We only retrain the model once due to the limited computation resources. Ideally, retraining the RL agent *each year*, which is aligned with the length of test period, leads the agent better utilized the up-to-date market

Table 2: Performance of different portfolio strategies with $\gamma = 3$

	Return	Volatility	Sharpe	Sortino	Max Drawdown	Calmer
S&P 500	10.98%	0.1460	0.7522	0.695519	-0.2477	0.4433
EW	13.26%	0.1502	0.8831	0.8310	-0.2293	0.5783
HistMarkowitz	8.08%	0.1553	0.5200	0.5226	-0.3014	0.2679
GenMarkowitz	13.59%	0.1536	0.8850	0.8526	-0.2690	0.5053
GenTD3	12.80%	0.1398	0.9158	0.8966	-0.2325	0.5507

Table 3: Performance of different portfolio strategies with $\gamma = 5$

	Return	Volatility	Sharpe	Sortino	Max Drawdown	Calmer
S&P 500	10.98%	0.1460	0.7522	0.6955	-0.2477	0.4433
EW	13.26%	0.1502	0.8831	0.8310	-0.2293	0.5783
HistMarkowitz	9.26%	0.1375	0.6737	0.6909	-0.2605	0.3555
GenMarkowitz	14.17%	0.1360	1.0418	1.0543	-0.1838	0.7711
GenTD3	12.73%	0.1392	0.9150	0.8916	-0.2326	0.5474

In Section 4 we present the portfolio performance when $\gamma = 0.5$. In this appendix we also present the same type of results with $\gamma = 3$ and $\gamma = 5$ to showcase the robustness.

We observe from Tables 2 and 3 that, although **GenMarkowitz** performs exceptional for $\gamma = 5$, it crashes for smaller $\gamma = 0.5$, even underperforming **EW** and S&P 500. **GenTD3**, in contrast, is stable in varying risk aversion and achieves the best results for $\gamma = 0.5$. It also has a performance comparable to **GenMarkowitz** for $\gamma = 5$, better than the benchmarks. This sensitivity analysis indicates the importance of the proposed adaptive sampling scheme. We also conclude that, the two strategies based on diffusion model are suitable for different risk aversions. In particular, choosing the appropriate risk aversion (which is not an easy task, though), **GenMarkowitz** may also have satisfactory performance, justifying the applicability of diffusion models in financial data.

E Algorithms

Algorithm 2 Training the score network and the RNN encoder

Initialization: Initialize s_θ and R_θ ; collect the training data.

- 1: **for** epochs **do**
- 2: Sample a batch of data $\{s_{(m)}^{1:T}\}_{m=1}^M$.
- 3: Compute $\{h_{(m)}^{1:T}\}_{m=1}^M$ using R_θ .
- 4: Sample $\{\tau_{(m)}\}_{m=1}^M$ from Uniform $[0, T]$.
- 5: Sample $\{z_{(m)}\}_{m=1}^M$ from $\mathcal{N}(0, I)$.
- 6: Update θ by minimizing (4).
- 7: **end for**

Outputs: Trained score network s_θ and RNN encoder R_θ .

F Proofs

Proof of Proposition B.1. By definition, it is not hard to show that

$$p_{t+1}(\tau, x|x^{1:t}) = \int_{\mathbb{R}^d} \phi(\tau, x|x_0) \mathbb{P}_{x^{1:t}}(dx_0).$$

model. However, this simple retraining scheme presented here has been sufficient to illustrate the application of our theories.

Algorithm 3 Adaptive sampling: implementation

Inputs: Trained score network s_θ and RNN encoder R_θ ; number time discretization (predictor) steps N_{pre} ; number of corrector steps N_{cor} ; predictor steps $\{\tau_k\}_{k=1}^{N_{\text{pre}}}$; corrector step sizes $\{\mathbf{e}_k\}_{k=1}^{N_{\text{pre}}}$; initial feature h^0

```
1:  $h \leftarrow h^0$ .
2: Sample  $s^{1:T} \sim \mathcal{N}(0_{\mathbb{R}^{dT}}, I_{dT \times dT})$ 
3: for  $t = 1, \dots, T$  do
4:   for  $k = 0, 1, \dots, N_{\text{pre}} - 1$  do
5:      $s^t \leftarrow s^t + \left( s^t + 2s_\theta(\mathcal{T} - \tau_k, h, s^t) \right) \cdot (\tau_{k+1} - \tau_k) + \sqrt{2}\varepsilon$ , where  $\varepsilon \sim \mathcal{N}(0, t_{k+1} - t_k)$ .
6:     # The predictor step.
7:     for  $l = 0, 1, \dots, N_{\text{cor}} - 1$  do
8:        $s^t \leftarrow s^t + \mathbf{e}_k s_\theta(\tau_k, h, s^t) + \sqrt{2}\varepsilon'$ , where  $\varepsilon' \sim \mathcal{N}(0, \mathbf{e}_k)$ .
9:       # The corrector step.
10:    end for
11:  end for
12:   $h \leftarrow R_\theta(s^t, h)$ . #Compute  $h^t$  from  $h^{t-1}$ .
13: end for
Outputs:  $s^{1:T}$ .
```

The term to be minimized in (7) can thus be calculated by

$$\begin{aligned} & \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) - \nabla_x \log p_{t+1}(\tau, X_\tau^{t+1} | X^{1:t}) \right|^2 \\ &= \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \right|^2 \\ & \quad + \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| \nabla_x \log p_{t+1}(\tau, X_\tau^{t+1} | X^{1:t}) \right|^2 \\ & \quad - 2\mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left[s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \cdot \nabla_x \log p_{t+1}(\tau, X_\tau^{t+1} | X^{1:t}) \right] \\ &= \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \right|^2 + C \\ & \quad - 2\mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \int_{\mathbb{R}^d} s_\theta^{t+1}(\tau, X^{1:t}, x) \cdot \nabla_x p_{t+1}(\tau, x | X^{1:t}) dx \\ &= \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \right|^2 + C \\ & \quad - 2\mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \int_{\mathbb{R}^d} s_\theta^{t+1}(\tau, X^{1:t}, x) \cdot \nabla_x \int_{\mathbb{R}^d} \phi(\tau, x | x_0) \mathbb{P}_{X^{1:t}}(dx_0) dx \\ &= \mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \right|^2 + C \\ & \quad - 2\mathbb{E}_{X^{1:t} \sim \mathbb{P}_{1:t}} \mathbb{E}_{X_\tau^{t+1} \sim \mathbb{P}_{X^{1:t}}} \mathbb{E}_{X_\tau^{t+1} \sim \phi(\tau, \cdot | X^{t+1})} \left[s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \cdot \nabla_x \log \phi(\tau, X_\tau^{t+1} | X^{t+1}) \right] \\ &= \mathbb{E}_{X^{1:t+1} \sim \mathbb{P}_{1:t+1}} \mathbb{E}_{X_\tau^{t+1} \sim p_{t+1}(\tau, \cdot | X^{1:t})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \right|^2 + C \\ & \quad - 2\mathbb{E}_{X^{1:t+1} \sim \mathbb{P}_{1:t+1}} \mathbb{E}_{X_\tau^{t+1} \sim \phi(\tau, \cdot | X^{t+1})} \left[s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) \cdot \nabla_x \log \phi(\tau, X_\tau^{t+1} | X^{t+1}) \right] \\ &= \mathbb{E}_{X^{1:t+1} \sim \mathbb{P}_{1:t+1}} \mathbb{E}_{X_\tau^{t+1} \sim \phi(\tau, \cdot | X^{t+1})} \left| s_\theta^{t+1}(\tau, X^{1:t}, X_\tau^{t+1}) - \nabla_x \log \phi(\tau, X_\tau^{t+1} | X^{t+1}) \right|^2 + C. \end{aligned}$$

Here, C is a generic constant independent from θ . The proof is completed because (7) and (8) attain the same minimizer. \square

To obtain the Wasserstein bounds from total variation bounds, we first provide a uniform-in-time moment estimate for SDE under *strong dissipative condition*; see Section 1.2 of Khasminskii [2012]. Suppose that for the function $f : [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^d$, there exists $M_{\text{disp}}, \delta > 0$ such that for any t, x ,

$$x \cdot f(t, x) \leq -(1 + \delta)|x|^2 + M_{\text{disp}}.$$

Algorithm 4 Policy gradient with generative environment

Initialization: Initialize Q-network Q_α and policy network π_β ; initialize target networks: $\alpha_{\text{targ}} \leftarrow \alpha$, $\beta_{\text{targ}} \leftarrow \beta$; set up an empty replay buffer \mathcal{D} and an empty scenario pool \mathcal{S} , both with size L ; train the score network s_θ and RNN encoder R_θ using Algorithm 2. Collect context window data $s^{-T:-1}$.

1: Use $s^{-T:-1}$ to compute initial feature h^0 .
2: Sample L paths prediction $s^{1:T}$ and compute corresponding feature paths $h^{1:T}$ using Algorithm 3; store them in \mathcal{S}

3: **for** steps **do**

4: Initialize $w = 1$; sample one path $(s^{1:T}, h^{1:T})$ from \mathcal{S} ; sample c from $\text{Exp}(10)$.

5: **for** $t = 1$ **to** $T - 1$ **do**

6: **if** warm-up **then**

7: $a \sim \text{Uniform}(\mathbb{K})$

8: **else**

9: $a \leftarrow \mathcal{P}_K(\pi_\beta(t, w, h^t, c) + \varepsilon)$.

10: **end if**

11: Store (w, a, c) in \mathcal{D} .

12: $w \leftarrow w + a \cdot (s^{t+1} - s^t)$.

13: **if not** warm-up **then**

14: Sample a batch $\{(s_{(j)}^{1:t+1}, h_{(j)}^{1:t+1})\}_{j=1}^B$ from \mathcal{S} and $\{(w_{(j)}, a_{(j)}, c_{(j)})\}_{j=1}^B$ from \mathcal{D} .

15: $w'_{(j)} = w_{(j)} + a_{(j)} \cdot (s_{(j)}^{t+1} - s_{(j)}^t)$.

16: **if** $t \leq T - 1$ **then**

17: Update α by minimizing

$$\frac{1}{B} \sum_j \left| Q_\alpha(t, w_{(j)}, h_{(j)}^t, a_{(j)}, c_{(j)}) - Q_{\alpha_{\text{targ}}}(t+1, w'_{(j)}, h_{(j)}^{t+1}, \pi_{\beta_{\text{targ}}}(t+1, w'_{(j)}, h_{(j)}^{t+1}, c_{(j)}), c_{(j)}) \right|^2.$$

18: **else**

19: Update α by minimizing

$$\frac{1}{B} \sum_j \left| Q_\alpha(t, w_{(j)}, h_{(j)}^t, a_{(j)}, c_{(j)}) - |w_{(j)} - c_{(j)}|^2 \right|^2.$$

20: **end if**

21: Update β by gradient ascent via

$$\frac{1}{B} \sum_j \nabla_a Q_\alpha(t, w_{(j)}, h_{(j)}^t, \pi_\beta(t, w_{(j)}, h_{(j)}^t, c_{(j)}), c_{(j)}) \nabla_\beta \pi_\beta(t, w_{(j)}, h_{(j)}^t, c_{(j)})$$

22: Update target network:

$$\alpha_{\text{targ}} \leftarrow \rho \alpha_{\text{targ}} + (1 - \rho) \alpha,$$

$$\beta_{\text{targ}} \leftarrow \rho \beta_{\text{targ}} + (1 - \rho) \beta.$$

23: **end if**

24: **end for**

25: **end for**

26: Find the optimal multiplier by maximizing

$$-\frac{\gamma}{2} Q_\alpha(1, 0, h^1, \pi_\beta(1, 0, h^1, c), c) + c.$$

Output: Policy network π_β

We consider the following SDE:

$$\begin{aligned} dX_t &= (X_t + f(t, X_t))dt + \sqrt{2}dB_t, \\ X_0 &\sim N(0, I). \end{aligned}$$

Lemma F.1. *There exist constants $c_1, c_2, c_3 > 0$, which are independent from t and M_{disp} , such that*

$$\sup_{t \geq 0} \mathbb{E}[e^{c_1|X_t|^2}] \leq c_2 e^{c_3 M_{\text{disp}}}. \quad (11)$$

Proof. Let $V(x) = e^{\theta|x|^2}$, where $\theta > 0$ will be determined later. Applying Itô's formula to $V(X_s)$, we have

$$\begin{aligned} V(X_t) - V(X_0) &= \int_0^t \left(2\theta \left(V(X_s) X_s \cdot (X_s + f(s, X_s)) \right) + 2\theta dV(X_s) + 4\theta^2 |X_s|^2 V(X_s) \right) ds \\ &\quad + 2\sqrt{2}\theta \int_0^t X_s V(X_s) dB_s. \end{aligned} \quad (12)$$

Considering a sequence of localization stopping time $\tau_n \rightarrow \infty$ and taking expectation on the above identity (12), we obtain

$$\begin{aligned} \mathbb{E}[V(X_{t \wedge \tau_n})] - \mathbb{E}[V(X_0)] &= \int_0^{t \wedge \tau_n} \left(2\theta \left(\mathbb{E}[V(X_s) X_s \cdot (X_s + f(s, X_s))] \right) \right. \\ &\quad \left. + 2\theta d\mathbb{E}[V(X_s)] + 4\theta^2 \mathbb{E}[|X_s|^2 V(X_s)] \right) ds \\ &= \int_0^t \left(2\theta \left(\mathbb{E}[V(X_{s \wedge \tau_n}) X_{s \wedge \tau_n} \cdot (X_{s \wedge \tau_n} + f(s \wedge \tau_n, X_{s \wedge \tau_n}))] \right) \right. \\ &\quad \left. + 2\theta d\mathbb{E}[V(X_{s \wedge \tau_n})] + 4\theta^2 \mathbb{E}[|X_{s \wedge \tau_n}|^2 V(X_{s \wedge \tau_n})] \right) ds \end{aligned}$$

The differential form yields

$$\begin{aligned} \frac{d}{dt} \mathbb{E}[V(X_{t \wedge \tau_n})] &= 2\theta \mathbb{E}[V(X_{s \wedge \tau_n}) X_{s \wedge \tau_n} \cdot (X_{s \wedge \tau_n} + f(s \wedge \tau_n, X_{s \wedge \tau_n}))] \\ &\quad + 2\theta d\mathbb{E}[V(X_{t \wedge \tau_n})] + 4\theta^2 \mathbb{E}[|X_{t \wedge \tau_n}|^2 V(X_{t \wedge \tau_n})] \\ &\leq (4\theta^2 - 2\theta\delta) \mathbb{E}[|X_{t \wedge \tau_n}|^2 V(X_{t \wedge \tau_n})] + 2\theta(M_{\text{disp}} + d) \mathbb{E}[V(X_{t \wedge \tau_n})]. \end{aligned}$$

Choose $\theta < \delta/2$, $R^2 > 2(M_{\text{disp}} + d)/(\delta - 2\theta)$ and denote $I_t = ((4\theta^2 - 2\theta\delta)|X_{t \wedge \tau_n}|^2 + 2\theta(M_{\text{disp}} + d))V(X_{t \wedge \tau_n})$. On the one hand, with $\gamma := 2\theta(M_{\text{disp}} + d)$, it is clear that $I_t \leq -\gamma V(X_{t \wedge \tau_n})$ when $|X_{t \wedge \tau_n}| \geq R$. On the other hand, when $|X_{t \wedge \tau_n}| \leq R$,

$$\begin{aligned} I_t &\leq 2\theta(M_{\text{disp}} + d)V(X_{t \wedge \tau_n}) \\ &\leq 2\theta(M_{\text{disp}} + d)V(R) \\ &= 2\theta(M_{\text{disp}} + d)e^{\frac{2\theta(M_{\text{disp}} + d)}{\delta - 2\theta}} \\ &\leq c'_1 e^{c'_2 M_{\text{disp}}}, \end{aligned}$$

where c'_1, c'_2 are some constants independent of t and M_{disp} . We thus obtain

$$\begin{aligned} \frac{d}{dt} \mathbb{E}[V(X_{t \wedge \tau_n})] &\leq -\gamma \mathbb{E}[V(X_{t \wedge \tau_n}) I_{|X_{t \wedge \tau_n}| > R}] + c'_1 e^{c'_2 M_{\text{disp}}} \\ &= -\gamma \mathbb{E}[V(X_{t \wedge \tau_n})] + \gamma \mathbb{E}[V(X_{t \wedge \tau_n}) I_{|X_{t \wedge \tau_n}| \leq R}] + c'_1 e^{c'_2 M_{\text{disp}}} \\ &\leq -\gamma \mathbb{E}[V(X_{t \wedge \tau_n})] + c'_3 e^{c'_4 M_{\text{disp}}} \end{aligned}$$

By Grownwall's inequality,

$$\mathbb{E}[V(X_{t \wedge \tau_n})] \leq e^{-\gamma t} \mathbb{E}[V(X_0) I_{|X_0| > R}] + \frac{c'_3 e^{c'_4 M_{\text{disp}}}}{\gamma} \leq c'_5 e^{c'_6 M_{\text{disp}}},$$

Here, constants $c'_1 \sim c'_6$ are all independent of t, n and M_{disp} . Thus, letting $n \rightarrow \infty$, (11) is established with $c_1 = \theta, c_2 = c'_5$, and $c_3 = c'_6$, all independent of t and M_{disp} . \square

Corollary F.2. For any $t, R > 0$, there exists $c_1, c_2 > 0$, which are independent of t and R , such that

$$\mathbb{E}[|X_t|^2 I_{\{|X_t| \geq R\}}] \leq e^{-c_1 R^2 + c_2 M_{\text{disp}}}.$$

Proof of Theorem 2.1. We first prove the following bound between conditional distributions:

There exists a constant $c > 0$ such that, for $t = 1, 2, \dots, T - 1$ we have

$$\mathcal{W}_2^2(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^{\mathcal{T}}) \leq C \left(\alpha(\mathcal{T}) + \mathcal{T}^2 \mathcal{E}(x^{1:t})^{1/2} + \mathcal{T}^{5/2} |x^{1:t} - y^{1:t}| \right), \quad (13)$$

where

$$\begin{aligned} \mathcal{E}(x^{1:t}) &:= \int_0^{\mathcal{T}} \mathbb{E}_{X_{\tau}^{t+1} \sim p_{t+1}(\tau, \cdot | x^{1:t})} |s_{\theta}^{t+1}(\tau, x^{1:t}, X_{\tau}^{t+1}) - \nabla_x \log p_{t+1}(\tau, X_{\tau}^{t+1} | x^{1:t})|^2 d\tau, \\ \alpha(\mathcal{T}) &:= \mathcal{T}^2 e^{-\mathcal{T}} + e^{-c\mathcal{T} + cM_{\text{disp}}}. \end{aligned}$$

To this end, consider the following two SDEs with initial conditions $\bar{X}_0 \sim p(\mathcal{T}, \cdot | x^{1:t})$ and $\bar{Y}_0 \sim \mathcal{N}(0, I)$,

$$\begin{aligned} d\bar{X}_{\tau} &= [\bar{X}_{\tau} + 2\nabla_x \log p_{t+1}(\mathcal{T} - \tau, \bar{X}_{\tau} | x^{1:t})] d\tau + \sqrt{2} dB_{\tau}, \\ d\bar{Y}_{\tau} &= [\bar{Y}_{\tau} + 2s_{\theta}^{t+1}(\mathcal{T} - \tau, y^{1:t}, \bar{Y}_{\tau})] d\tau + \sqrt{2} dB_{\tau}. \end{aligned}$$

By the proof of total variation bounds for continuous time DDPM (see, e.g., Song et al. [2021a] and Chen et al. [2023]), we have

$$\begin{aligned} \text{TV}(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^{\mathcal{T}})^2 &\leq C \left(e^{-2\mathcal{T}} + \int_0^{\mathcal{T}} \mathbb{E}_{X_{\tau}^{t+1} \sim p_{t+1}(\tau, \cdot | x^{1:t})} |s_{\theta}^{t+1}(\tau, y^{1:t}, X_{\tau}^{t+1}) - \nabla_x \log p_{t+1}(\tau, X_{\tau}^{t+1} | x^{1:t})|^2 d\tau \right) \\ &\leq C(e^{-2\mathcal{T}} + \mathcal{E}(x^{1:t}) + \mathcal{T}|x^{1:t} - y^{1:t}|^2) \end{aligned}$$

Therefore, recalling that $\sqrt{a+b+c} \leq \sqrt{a} + \sqrt{b} + \sqrt{c}$ for $a, b, c \geq 0$, we have

$$\text{TV}(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^{\mathcal{T}}) \leq C(e^{-\mathcal{T}} + \mathcal{E}(x^{1:t})^{1/2} + \mathcal{T}^{1/2}|x^{1:t} - y^{1:t}|). \quad (14)$$

To bound $\mathcal{W}(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}})$ with their total variation metric, we use the following classical results regarding these two different metrics (see Villani [2008], Theorem 6.15): for any two probability measures μ, ν and any $R > 0$,

$$\mathcal{W}_2^2(\mu, \nu) \leq R^2 \text{TV}(\mu, \nu) + \int_{|x| \geq R} |x|^2 (\mu + \nu)(dx).$$

Taking $\mu = \mathbb{P}_{x^{1:t}}, \nu = \mathbb{Q}_{y^{1:t}}$, we have

$$\mathcal{W}_2^2(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^{\mathcal{T}}) \leq C \left(R^2 \text{TV}(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^{\mathcal{T}}) + \mathbb{E}_{\mathbb{P}_{x^{1:t}}} [|X^{t+1}|^2 I_{\{|X^{t+1}| \geq R\}}] + \mathbb{E}_{\mathbb{Q}_{y^{1:t}}^{\mathcal{T}}} [|Y^{t+1}|^2 I_{\{|Y^{t+1}| \geq R\}}] \right). \quad (15)$$

We now estimate three terms on the right-hand side of (15). By Assumption 2,

$$\begin{aligned} \mathbb{E}_{\mathbb{P}_{x^{1:t}}} [|X^{t+1}|^2 I_{\{|X^{t+1}| \geq R\}}] &\leq C \mathbb{E}_{\mathbb{P}_{x^{1:t}}} [e^{\frac{c}{2}|X^{t+1}|} I_{\{|X^{t+1}| \geq R\}}] \\ &\leq C \left\{ \mathbb{E}_{\mathbb{P}_{x^{1:t}}} [e^{c|X^{t+1}|}] \right\}^{1/2} \left\{ \mathbb{P}_{x^{1:t}} (|X^{t+1}| \geq R) \right\}^{1/2} \\ &\leq C e^{-\frac{c}{2}R} \mathbb{E}_{\mathbb{P}_{x^{1:t}}} [e^{c|X^{t+1}|}]. \\ &\leq C e^{-\frac{c}{2}R} \end{aligned} \quad (16)$$

By Corollary F.2,

$$\begin{aligned} \mathbb{E}_{\mathbb{Q}_{y^{1:t}}^{\mathcal{T}}} [|Y^{t+1}|^2 I_{\{|Y^{t+1}| \geq R\}}] &= \mathbb{E}_{\mathbb{Q}^{\mathcal{T}}} [|\bar{Y}_{\mathcal{T}}|^2 I_{\{|\bar{Y}_{\mathcal{T}}| \geq R\}}] \\ &\leq e^{-c_1 R^2 + c_2 M_{\text{disp}}}. \end{aligned} \quad (17)$$

Plugging (14), (16) and (17) into (15), we have

$$\begin{aligned} \mathcal{W}_2(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^T)^2 &\leq C(R^2 e^{-\mathcal{T}} + R^2 \mathcal{E}(x^{1:t})^{1/2} + R^2 \mathcal{T}^{1/2} |x^{1:t} - y^{1:t}| + C e^{-\frac{\varepsilon}{2} R} + e^{-c_1 R^2 + c_2 M_{\text{disp}}}) \\ &\leq C(R^2 e^{-\mathcal{T}} + R^2 \mathcal{E}(x^{1:t})^{1/2} + R^2 \mathcal{T}^{1/2} |x^{1:t} - y^{1:t}| + e^{-cR + cM_{\text{disp}}}). \end{aligned}$$

Taking $R = \mathcal{T}$, we conclude (13).

Next, we prove the main result (2). Because $\pi \in \Pi_{\text{bc}}(\mathbb{P}, \mathbb{Q}^T)$ if and only if $\pi_{x^{1:t}, y^{1:t}} \in \Pi(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^T)$ (Backhoff et al. [2017], Proposition 5.1), we construct a specific coupling $\pi^\varepsilon \in \Pi_{\text{bc}}(\mathbb{P}, \mathbb{Q})$ as follows:

- (a) $\int_{\mathbb{R}^d \times \mathbb{R}^d} |x^1 - y^1|^2 \pi_1^\varepsilon(dx^1, dy^1) \leq \mathcal{W}_2^2(\mathbb{P}_1, \mathbb{Q}_1^T) + \varepsilon;$
- (b) For $t = 1, 2, \dots, T - 1$, $\int_{\mathbb{R}^d \times \mathbb{R}^d} |x^{t+1} - y^{t+1}|^2 \pi_{x^{1:t}, y^{1:t}}^\varepsilon(dx^{t+1}, dy^{t+1}) \leq \mathcal{W}_2^2(\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}^T) + \varepsilon;$
- (c) $\pi^\varepsilon(dx^{1:T}, dy^{1:T}) := \pi_1^\varepsilon(dx^1, dy^1) \pi_{x^1, y^1}^\varepsilon(dx^2, dy^2) \cdots \pi_{x^{1:T-1}, y^{1:T-1}}^\varepsilon(dx^T, dy^T).$

We first use induction to show that, for $t = 1, 2, \dots, T$,

$$\mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 \leq C \left(\mathcal{T}^{\frac{5t}{2}} \varepsilon_{\text{score}}^{1/2^{t-1}} + \mathcal{T}^{\frac{5(t-1)}{2}} (\alpha(\mathcal{T}) + \varepsilon)^{1/2^{t-1}} \right). \quad (18)$$

By repeating the proof of (1), with $\mathbb{P}_{x^{1:t}}, \mathbb{Q}_{y^{1:t}}$ replaced by $\mathbb{P}_1, \mathbb{Q}_1$, we can prove that (18) holds for $t = 1$. This is because when considering non-conditional probability \mathbb{P}_1 and \mathbb{Q}_1 , the term $|x^{1:t} - y^{1:t}|$ disappears and the term $\mathcal{E}(x^{1:t})$ is upper bounded by $\mathcal{T} \varepsilon_{\text{score}}^2$.

Suppose now that (18) is true for t , we aim to prove it for $t + 1$. Indeed, by using (13), we have

$$\begin{aligned} \mathbb{E}_{\pi_{1:t+1}^\varepsilon} [|X^{1:t+1} - Y^{1:t+1}|^2] &\leq C \left(\mathbb{E}_{\pi_{1:t}^\varepsilon} [|X^{1:t} - Y^{1:t}|^2] + \mathbb{E}_{\pi_{1:t}^\varepsilon} \mathbb{E}_{\pi_{x^{1:t}, y^{1:t}}^\varepsilon} [|X^{t+1} - Y^{t+1}|^2] \right) \\ &\leq C \left(\mathbb{E}_{\pi_{1:t}^\varepsilon} [|X^{1:t} - Y^{1:t}|^2] + \mathbb{E}_{\pi_{1:t}^\varepsilon} [\mathcal{W}_2^2(\mathbb{P}_{X^{1:t}}, \mathbb{Q}_{Y^{1:t}}^T)] + \varepsilon \right) \\ &\leq C \left(\mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 + \mathcal{T}^2 \mathbb{E}_{\pi_{1:t}^\varepsilon} \mathcal{E}(X^{1:t})^{1/2} + \mathcal{T}^{\frac{5}{2}} \{ \mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 \}^{1/2} \right. \\ &\quad \left. + \alpha(\mathcal{T}) + \varepsilon \right) \end{aligned} \quad (19)$$

With (6), we have

$$\mathbb{E}_{\pi_{1:t}^\varepsilon} [\mathcal{E}(X^{1:t})^{1/2}] = \mathbb{E}_{\mathbb{P}_{1:t}} [\mathcal{E}(X^{1:t})^{1/2}] \leq \{ \mathbb{E}_{\mathbb{P}_{1:t}} [\mathcal{E}(X^{1:t})] \}^{1/2} \leq \mathcal{T}^{1/2} \varepsilon_{\text{score}}.$$

If $\mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 > 1$, then from (19) we know

$$\begin{aligned} \mathbb{E}_{\pi_{1:t+1}^\varepsilon} |X^{1:t+1} - Y^{1:t+1}|^2 &\leq C \left(\mathcal{T}^{\frac{5}{2}} \mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 + \mathcal{T}^{\frac{5}{2}} \varepsilon_{\text{score}} + \alpha(\mathcal{T}) + \varepsilon \right) \\ &\leq C \left(\mathcal{T}^{\frac{5(t+1)}{2}} \varepsilon_{\text{score}}^{1/2^{t-1}} + \mathcal{T}^{\frac{5t}{2}} (\alpha(\mathcal{T}) + \varepsilon)^{1/2^{t-1}} \right. \\ &\quad \left. + \mathcal{T}^{\frac{5}{2}} \varepsilon_{\text{score}} + \alpha(\mathcal{T}) + \varepsilon \right) \\ &\leq C \left(\mathcal{T}^{\frac{5(t+1)}{2}} \varepsilon_{\text{score}}^{1/2^t} + \mathcal{T}^{\frac{5t}{2}} (\alpha(\mathcal{T}) + \varepsilon)^{1/2^t} \right). \end{aligned}$$

Here, we use $\mathcal{T} > 1$, $\varepsilon_{\text{score}} < 1$ and choose a small ε such that $\alpha(\mathcal{T}) + \varepsilon < 1$. Similarly, if $\mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 \leq 1$,

$$\begin{aligned} \mathbb{E}_{\pi_{1:t+1}^\varepsilon} |X^{1:t+1} - Y^{1:t+1}|^2 &\leq C \left(\mathcal{T}^{\frac{5}{2}} \{ \mathbb{E}_{\pi_{1:t}^\varepsilon} |X^{1:t} - Y^{1:t}|^2 \}^{1/2} + \mathcal{T}^{\frac{5}{2}} \varepsilon_{\text{score}} + \alpha(\mathcal{T}) + \varepsilon \right) \\ &\leq C \left(\mathcal{T}^{\frac{5}{2}} \left(\frac{t}{2} + 1 \right) \varepsilon_{\text{score}}^{1/2^t} + \mathcal{T}^{\frac{5}{2}} \left(\frac{t-1}{2} + 1 \right) (\alpha(\mathcal{T}) + \varepsilon)^{1/2^t} \right. \\ &\quad \left. + \mathcal{T}^{\frac{5}{2}} \varepsilon_{\text{score}} + \alpha(\mathcal{T}) + \varepsilon \right) \\ &\leq C \left(\mathcal{T}^{\frac{5(t+1)}{2}} \varepsilon_{\text{score}}^{1/2^t} + \mathcal{T}^{\frac{5t}{2}} (\alpha(\mathcal{T}) + \varepsilon)^{1/2^t} \right). \end{aligned}$$

In either case, we establish that (18) holds for $t + 1$, hence for all $t \in \{1, 2, \dots, T\}$. Taking $t = T$ in (18) and using the definition of adapted Wasserstein metric, we obtain

$$\begin{aligned} \mathcal{AW}_2^2(\mathbb{P}, \mathbb{Q}^T) &\leq \mathbb{E}_{\pi^\varepsilon} |X^{1:T} - Y^{1:T}|^2 \\ &\leq C \left(\mathcal{T}^{\frac{5T}{2}} \varepsilon_{\text{score}}^{1/2^{T-1}} + \mathcal{T}^{\frac{5(T-1)}{2}} (\alpha(\mathcal{T}) + \varepsilon)^{1/2^{T-1}} \right). \end{aligned} \quad (20)$$

Letting $\varepsilon \rightarrow 0$ in (20) gives

$$\begin{aligned} \mathcal{AW}_2^2(\mathbb{P}, \mathbb{Q}^T) &\leq \mathbb{E}_{\pi^\varepsilon} |X^{1:T} - Y^{1:T}|^2 \\ &\leq C \left(\mathcal{T}^{\frac{5T}{2}} \varepsilon_{\text{score}}^{1/2^{T-1}} + \mathcal{T}^{\frac{5(T-1)}{2}} (\mathcal{T}^2 e^{-\mathcal{T}} + e^{-c\mathcal{T} + cM_{\text{disp}}})^{1/2^{T-1}} \right). \end{aligned}$$

Using Proposition C.1 we take $M_{\text{disp}} \sim \log(1/\varepsilon_{\text{score}})$ and obtain (2). \square

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- Please refer to our LLM policy (<https://neurips.cc/Conferences/2025/LLM>) for what should or should not be described.