Bitrate-Constrained DRO: Beyond Worst Case Robustness To Unknown Group Shifts

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Abstract

Although training machine learning models for robustness is critical for real-world 1 2 adoption, determining how to best ensure robustness remains an open problem. 3 Some methods (e.g., DRO) are overly conservative, while others (e.g., Group DRO) require domain knowledge that may be hard to obtain. In this work, we address 4 limitations in prior approaches by assuming a more nuanced form of group shift: 5 conditioned on the label, we assume that the true group function is *simple*. For 6 example, we may expect that group shifts occur along high-level features (e.g., 7 image background, lighting). Thus, we aim to learn a model that maintains high 8 accuracy on simple group functions realized by these features, but need not spend 9 valuable model capacity achieving high accuracy on contrived groups of examples. 10 Based on this idea, we formulate a two-player game where conditioned on the 11 label the adversary can only separate datapoints into potential groups using simple 12 features, which corresponds to a bitrate constraint on the adversary's capacity. Our 13 resulting practical algorithm, Bitrate-Constrained DRO (BR-DRO), does not require 14 group information on training samples yet matches the performance of Group DRO. 15 Our theoretical analysis reveals that in some settings BR-DRO objective can provably 16 yield statistically efficient and less pessimistic solutions than unconstrained DRO. 17

18 1 Introduction

A common form of distribution shift is group shift, where the source and target differ only in the 19 marginal distribution over finite groups or sub-populations, with no change in group conditionals [43, 20 18, 46]. Prior works consider various approaches to address group shift. One solution is to ensure 21 robustness to worst case shifts using distributionally robust optimization (DRO) [4, 7, 17], which 22 considers a two-player game where a learner minimizes risk on distributions chosen by an adversary 23 from a predefined uncertainty set. As the adversary is unconstrained in proposing distributions, 24 DRO often yields overly pessimistic solutions [25] and can suffer from statistical challenges [18]. 25 Methods like Group DRO [46] avoid overly pessimistic solutions by assuming knowledge of group 26 membership for each training example. However, these group-based methods provide no guarantees 27 on shifts that deviate from the predefined groups, and are not applicable to problems that lack group 28 knowledge. In this work, we therefore ask: Can we train non-pessimistic robust models without 29 access to group annotations on training examples? 30

We address this question by considering a more nuanced assumption on the structure of the underlying groups. We assume that, conditioned on the label, group boundaries are realized by high-level features that depend on a small set of underlying factors. This leads to simpler group functions with large margins and simple decision boundaries (Figure 1 *(left)*). Invoking the principle of minimum description length [21], restricting our adversary to functions that satisfy this assumption corresponds to a bitrate constraint. In DRO, the adversary upweights points with higher losses under the current

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Figure 1: **Bitrate-Constrained DRO**: A method that assumes group shifts along low-bitrate features, and restricts the adversary appropriately so that the solution found is less pessimistic and more robust to group shifts. Our method is also robust to training noise. (*Left*) In Waterbirds [54], the spurious feature background is a large margin simple feature that separates the *majority* and *minority* points in each class. (*Right*) Prior works [31, 35] that upweight arbitrary points with high losses force the model to memorize noisy mislabeled points while our method is robust to noise and only upweights the true minority group without any knowledge of its identity.

³⁷ learner, which in practice often correspond to examples that belong to a rare group, contain complex

patterns, or are mislabeled [14, 53]. Restricting the adversary's capacity prevents it from upweighting

individual hard or mislabeled examples (as they cannot be identified with simple features), and biases

40 it towards identifying erroneous data points misclassified by simple features. This also complements

- the failure mode of neural networks trained with stochastic gradient descent (SGD) that rely on simple spurious features that correctly classify points in the *majority* group but may fail on *minority*
- 43 groups [10].

The main contribution of this paper is Bitrate-Constrained DRO (BR-DRO), a supervised learning 44 procedure that provides robustness to distribution shifts along groups realized by simple functions. 45 Despite not using group information on training examples, we demonstrate that BR-DRO can match the 46 performance of methods requiring them. We also find that BR-DRO correctly identifies true minority 47 48 points, whereas DRO without group information does not. This indicates that not optimizing for performance on contrived worst-case shifts can reduce the pessimism inherent in DRO. It further 49 validates: (i) our assumption on the simple nature of group shift; and (ii) that our method of capacity 50 control meaningfully structures the uncertainty set to be robust to such shifts. As a consequence of 51 the constraint, we also find that BR-DRO is robust to random noise in the training data [51], since 52 it cannot form "groups" entirely based on randomly mislabeled points with low bitrate features. 53 This is in contrast with existing methods that use the learner's training error to up-weight arbitrary 54 sets of difficult training points [e.g., 35, 31], which we show are highly susceptible to label noise 55 (see Figure 1). Finally, we theoretically analyze our approach-characterizing how the degree of 56 constraint on the adversary can effect worst risk estimation and excess risk (pessimism) bounds, as 57 58 well as convergence rates for specific online solvers.

59 2 Bitrate-Constrained DRO

Notation. With covariates $\mathcal{X} \subset \mathbb{R}^d$ and labels \mathcal{Y} , the given source P and unknown true target Q_0 are measures over the measurable space $(\mathcal{X} \times \mathcal{Y}, \Sigma)$ and have densities p and q_0 respectively (w.r.t. base measure μ). The learner's choice is a hypothesis $h : \mathcal{X} \mapsto \mathcal{Y}$ in class $\mathcal{H} \subset L^2(P)$, and the adversary's action in standard DRO is a target distribution Q in set $Q_{P,\kappa} := \{Q : Q \ll P, D_f(Q || P) \le \kappa\}$. Here, D_f is the f-divergence between Q and P for a convex function f^1 with f(1) = 0. An

equivalent action space for the adversary is the set of re-weighting functions:

$$\mathcal{W}_{P,\kappa} = \{ w : \mathcal{X} \times \mathcal{Y} \mapsto \mathbb{R} : w \text{ is measurable under } P, \ \mathbb{E}_P[w] = 1, \ \mathbb{E}_P f(w) \le \kappa \}$$
(1)

For a convex loss function $l : \mathcal{Y} \times \mathcal{Y} \mapsto \mathbb{R}_+$, we denote l(h) as the function over (\mathbf{x}, \mathbf{y}) that evaluates $l(h(\mathbf{x}), \mathbf{y})$, and use l_{0-1} to denote the loss function $\mathbb{I}(h(\mathbf{x}) \neq \mathbf{y})$. Given either distribution $Q \in \mathcal{Q}_{P,\kappa}$, or a re-weighting function $w \in \mathcal{W}_{P,\kappa}$, the risk of a learner h is:

$$R(h,Q) = \mathbb{E}_Q \left[l(h) \right] \qquad R(h,w) = \mathbb{E}_{(\mathbf{x},\mathbf{y})\sim P} \left[l(h(\mathbf{x}),\mathbf{y}) \cdot w(\mathbf{x},\mathbf{y}) \right] = \langle l(h), w \rangle_P \qquad (2)$$

Note the overload of notation for $R(h, \cdot)$. If the adversary is stochastic it picks a mixed action

 $\delta \in \Delta(\mathcal{W}_{P,\kappa})$, which is the set of all distributions over $\mathcal{W}_{P,\kappa}$. Whenever it is clear, we drop P, κ .

71 Unconstrained DRO [7]. This is a min-max optimization problem understood as a two-player game,

⁷² where the learner chooses a hypothesis, to minimize risk on the worst distribution that the adversary

⁷³ can choose from its set. Formally, this is given by equation 3. The first equivalence is clear from the

⁷⁴ definitions and for the second since R(h, Q) is linear in Q, the supremum over $\Delta(W_{P,\kappa})$ is a Dirac

¹For e.g., KL(Q || P) can be derived with $f(x) = x \log x$ and for Total Variation f(x) = |x - 1|/2.

⁷⁵ delta over the best weighting in $W_{P,\kappa}$. In the next section, we will see how a bitrate-constrained ⁷⁶ adversary can only pick certain actions from $\Delta(W_{P,\kappa})$.

$$\inf_{h \in \mathcal{H}} \sup_{Q \in \mathcal{Q}_{P,\kappa}} R(h,Q) \equiv \inf_{h \in \mathcal{H}} \sup_{w \in \mathcal{W}_{P,\kappa}} R(h,w) \equiv \inf_{h \in \mathcal{H}} \sup_{\delta \in \Delta(\mathcal{W}_{P,\kappa})} \mathbb{E}_{w \sim \delta} \left[R(h,w) \right]$$
(3)

Group Shift. While the DRO framework is broad and addresses any unstructured shift, we focus on the specific case of group shift. First, for a given pair of measures P, Q we define what we mean by the group structure $\mathcal{G}_{P,Q}$ (Definition 2.1). Intuitively, it is a set of sub-populations along which the distribution shifts, defined in a way that makes them uniquely identifiable. For *e.g.*, in the Waterbirds dataset (Figure 1), there are four groups given by combinations of (label, background). Corollary 2.2 follows immediately from the definition of $\mathcal{G}_{P,Q}$. Using this definition, the standard group shift assumption [46] can be formally re-stated as Assumption 2.3.

Definition 2.1 (group structure $\mathcal{G}_{P,Q}$). For $Q \ll P$ the group structure $\mathcal{G}_{P,Q} = \{G_k\}_{k=1}^K$ is the smallest finite set of disjoint groups $\{G_k\}_{k=1}^K$ s.t. $Q(\cup_{k=1}^K G_k) = 1$ and $\forall k$ (i) $G_k \in \Sigma$, $Q(G_k) > 0$ and (ii) $p(\mathbf{x}, \mathbf{y} \mid G_k) = q(\mathbf{x}, \mathbf{y} \mid G_k) > 0$ a.e. in μ . If such a structure exists then $\mathcal{G}_{P,Q}$ is well defined.

Corollary 2.2 (uniqueness of $\mathcal{G}_{P,Q}$). $\forall P, Q$, the structure $\mathcal{G}(P,Q)$ is unique if it is well defined.

Assumption 2.3 (standard group shift). There exists a well-defined group structure \mathcal{G}_{P,Q_0} s.t. target \mathcal{Q}_{Q_0} differs from P only in terms of maximal probabilities over all $C \in \mathcal{G}_{P,Q_0}$

⁹⁰ Q_0 differs from P only in terms of marginal probabilities over all $G \in \mathcal{G}_{P,Q_0}$.

How expressive is unconstrained adversary? Note that the set $W_{P,\kappa}$ includes all measurable 91 functions (under P) such that the re-weighted distribution is bounded in f-divergence (by κ). While 92 prior works [48, 17] shrink κ to construct confidence intervals, this only controls the total mass that 93 can be moved between measurable sets $G_1, G_2 \in \Sigma$, but *does not restrict* the choice of G_1 and G_2 94 itself. As noted by Hu et al. [25], such an adversary is highly expressive, and optimizing for the 95 worst case only leads to the solution of empirical risk minimization (ERM) under l_{0-1} loss. Thus, 96 we can conclude that DRO recovers degenerate solutions because the worst target in $W_{P,\kappa}$ lies far 97 from the subspace of naturally occurring targets. Since it is hard to precisely characterize natural 98 targets we make a nuanced assumption: the target Q_0 only upsamples those rare subpopulations that 99 are misclassified by simple features. We state this formally in Assumption 2.5 after we define the 100 bitrate-constrained function class $\mathcal{W}(\gamma)$ in Definition 2.4. See Appendix A for additional discussion 101 on when/why constraining capacity helps with distribution shift robustness. 102 **Definition 2.4.** A function class $\mathcal{W}(\gamma)$ is bitrate-constrained if there exists a data independent prior 103

104 π , s.t. $\mathcal{W}(\gamma) = \{\mathbb{E}[\delta] : \delta \in \Delta(\mathcal{W}), \ \textit{KL}(\delta \mid\mid \pi) \leq \gamma\}.$

Assumption 2.5 (simple group shift). Target Q_0 satisfies Assumption 2.3 (group shift) w.r.t. source

106 P. Additionally, for some prior π and a small γ^* , the re-weighting function q_0/p lies in a bitrate-107 constrained class $\mathcal{W}(\gamma^*)$. In other words, for every group $G \in \mathcal{G}(P,Q_0)$, $\exists w_G \in \mathcal{W}(\gamma^*)$ s.t.

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$$\mathbb{1}((\mathbf{x}, \mathbf{y}) \in G) = w_G a.e.$$
 We refer to such a G as a simple group that is realized in $\mathcal{W}(\gamma^*)$.

BR-DRO **objective.** According to Assumption 2.5, there cannot exist a target Q_0 such that minority $G_{\min} \in \mathcal{G}(P, Q_0)$ is not realized in bitrate constrained class $\mathcal{W}(\gamma^*)$. Thus, by constraining our adversary to a class $\mathcal{W}(\gamma)$ (for some γ that is user defined), we can possibly evade issues emerging from optimizing for performance on mislabeled or hard examples, even if they were rare. This gives us the objective in Equation 4 where the equalities hold from the linearity of $\langle \cdot, \cdot \rangle$ and Definition 2.4. See Appendix A.1 for details on the practical implementation of BR-DRO.

$$\inf_{h \in \mathcal{H}} \sup_{\substack{\delta \in \Delta(\mathcal{W}) \\ \operatorname{KL}(\delta \mid\mid \pi) < \gamma}} \mathbb{E}_{w \sim \delta} R(h, w) = \inf_{h \in \mathcal{H}} \sup_{\substack{\delta \in \Delta(\mathcal{W}) \\ \operatorname{KL}(\delta \mid\mid \pi) < \gamma}} \langle l(h), \mathbb{E}_{\delta}[w] \rangle_{P} = \inf_{h \in \mathcal{H}} \sup_{w \in \mathcal{W}(\gamma)} R(h, w)$$
(4)

Theoretical Analysis. The main objective of our analysis of BR-DRO is to show how adding a bitrate constraint on the adversary can: (i) give us tighter statistical estimates of the worst risk; and (ii) control the pessimism (excess risk) of the learned solution. First, we provide worst risk generalization guarantees using the PAC-Bayes framework [15], along with a result for kernel adversary. Then, we discuss convergence rates and pessimism guarantees for the solution found by our online solver for a specific instance of $W(\gamma)$. See Appendix B for details.

121 **3 Experiments**

We discuss two sets of experiments here on robustness to spurious correlations and random label noise. For more details on these and other experiments please refer to Appendix C.

	Waterbirds		Cel	ebA	CivilComments	
Method	Avg	WG	Avg	WG	Avg	WG
ERM	97.1 (0.1)	71.0 (0.4)	95.4 (0.2)	46.9 (1.0)	92.3 (0.2)	57.2 (0.9)
LfF [41]	90.7 (0.2)	77.6 (0.5)	85.3 (0.2)	77.4 (0.7)	92.4 (0.1)	58.9 (1.1)
RWY [26]	93.7 (0.3)	85.8 (0.5)	84.9 (0.2)	80.4 (0.3)	91.7 (0.2)	67.7 (0.7)
JTT [35]	93.2 (0.2)	86.6 (0.4)	87.6 (0.2)	81.3 (0.5)	90.8 (0.3)	69.4 (0.8)
CVaR DRO [31]	96.3 (0.2)	75.5 (0.4)	82.2 (0.3)	64.7 (0.6)	92.3 (0.2)	60.2 (0.8)
BR-DRO (VIB) (ours)	94.1 (0.2)	86.3 (0.3)	86.7 (0.2)	80.9 (0.4)	90.5 (0.2)	68.7 (0.9)
BR-DRO (l_2) (ours)	93.8 (0.2)	86.4 (0.3)	87.7 (0.3)	80.4 (0.6)	91.0 (0.3)	68.9 (0.7)
Group DRO [46]	93.2 (0.3)	91.1 (0.3)	92.3 (0.3)	88.4 (0.6)	88.5 (0.3)	70.0 (0.5)

Table 1: BR-DRO recovers worst group performance gap between CVaR DRO and Group DRO: On Waterbirds, CelebA and CivilComments we report test average (Avg) and test worst group (WG) accuracies for BR-DRO and baselines. In (\cdot) we report the standard error of the mean accuracy across five runs.



Figure 2: (*Left*) Visualization (2d) of noisy synthetic data and learned predictors: We compare the decision boundaries (projected onto core and spurious features) learned by JTT with BR-DRO when the adversary is restricted to a sparse predictor. While our method recovers the core feature the baselines memorize the minority points. (*Right*) BR-DRO is robust to random label noise in training data: Across varying levels of the fraction of noise in training data we compare performance of BR-DRO with ERM and methods (JTT, CVaR DRO) that naively up weight high loss datapoints.

Is BR-DRO robust to group shifts without training data group annotations? Table 1 compares the 124 average and worst group accuracy for BR-DRO with ERM and four group shift robustness baselines: 125 JTT, LtF, SUBY, and CVaR DRO. First, we see that unconstrained CVaR DRO underperforms other 126 heuristic algorithms. This matches the observation made by Liu et al. [35]. Next, we see that adding 127 bitrate constraints on the adversary via a KL term or l_2 penalty significantly improves the performance 128 of BR-DRO (VIB) or BR-DRO (l_2) , which now matches the best performing baseline (JTT). Thus, we 129 see the less conservative nature of BR-DRO allows it to recover a large portion of the performance 130 gap between Group DRO and CVaR DRO. Indirectly, this partially validates our Assumption 2.5, 131 which states that the minority group is identified by a low bitrate adversary class. In Section C.3 we 132 discuss exactly what fraction of the minority group is identified, and the role played by the strength 133 of bitrate-constraint. 134

Bitrate DRO is more robust to random label noise. Several methods for group robustness (e.g., 135 CVaR DRO, JTT) are based on the idea of up weighting points with high training losses. The goal 136 is to obtain a learner with matching performance on every (small) fraction of points in the dataset. 137 However, when training data has mislabeled examples, such an approach will likely yield degenerate 138 solutions. This is because the adversary directly upweights any example where the learner has high 139 loss, including datapoints with incorrect labels. Hence, even if the learner's prediction matches the 140 (unknown) true label, this formulation would force the learner to memorize incorrect labelings at the 141 expense of learning the true underlying function. On the other hand, if the adversary is sufficiently 142 bitrate constrained, it cannot upweight the arbitrary set of randomly mislabeled points, as this would 143 require it to memorize those points. Our Assumption 2.5 also dictates that the distribution shift would 144 not upsample such high bitrate noisy examples. Thus, our constraint on the adversary ensures BR-DRO 145 is robust to label noise in the training data and our assumption on the target distribution retains its 146 robustness to test time distribution shifts. In Figure 2b we highlight this failure mode of unconstrained 147 up-weighting methods in contrast to BR-DRO. We first induce random label noise [14] of varying 148 degrees into the Waterbirds and CelebA training sets. Then we run each method and compare 149 worst group performance. In the presence of noise, BR-DRO significantly outperforms JTT and other 150 approaches on both Waterbirds and CelebA, as it only upsamples the minority examples misclassified 151 by simple features, ignoring the noisy examples for the reasons above. See Appendix C.1 for more 152 details on experiments with synthetic data. 153

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295 Appendix

296 A Additional discussion on Bitrate-Constrained DRO

Note on Assumption 2.5. Under the principle of minimum description length [21] any deviation from the prior (*i.e.*, KL($\delta \mid \mid \pi$)) increases the *description length* of the encoding $\delta \in \Delta(W)$, thus we refer to $W(\gamma)$ as being *bitrate-constrained* in the sense that it contains functions (means of distributions) that can be described with a limited number of bits given the prior π . Next we present arguments for why identifiability of simple (satisfy Assumption 2.5) minority groups can be critical for robustness.

Neural networks can perform poorly on simple minorities. For a fixed target Q_0 , let's say there 303 exists two groups: G_{\min} and $G_{\max} \in \mathcal{G}(P, Q_0)$ such that $P(G_{\min}) \ll P(G_{\max})$. By Assumption 2.5, 304 both G_{\min} and G_{\max} are simple (realized in $\mathcal{W}(\gamma^*)$), and are thus separated by some simple feature. 305 The learner's class \mathcal{H} is usually a class of overparameterized neural networks. When trained with 306 stochastic gradient descent (SGD), these are biased towards learning simple features that classify a 307 majority of the data [49, 52]. Thus, if the simple feature separating G_{\min} and G_{\max} itself correlates with the label y on G_{\max} , then neural networks would fit on this feature. This is precisely the case 308 309 in the Waterbirds example, where the groups are defined by whether the simple feature background 310 correlates with the label (Figure 1). Thus our assumption on the nature of shift complements the 311 nature of neural networks perform poorly on simple minorities. 312

The bitrate constraint helps identify simple unfair minorities in $\mathcal{G}(P, Q_0)$. Any method that aims 313 to be robust on Q_0 must up-weight data points from G_{\min} but without knowing its identity. Since 314 the unconstrained adversary upsamples any group of data points with high loss and low probability, 315 it cannot distinguish between a rare group that is realized by simple functions in $\mathcal{W}(\gamma^*)$ and a rare 316 group of examples that share no feature in common or may even be mislabeled. On the other hand, 317 the group of mislabeled examples cannot be separated from the rest by functions in $\mathcal{W}(\gamma^*)$. Thus, 318 a bitrate constraint adversary can only identify simple groups and upsamples those that incur high 319 losses – possibly due to the simplicity bias of neural networks. 320

321 A.1 Bitrate-Constrained DRO in Practice

BR-DRO in practice. We parameterize the learner $\theta_h \in \Theta_h$ and adversary $\theta_w \in \Theta_w$ as neural 322 networks². Therefore, the objective in Equation 5 is no longer convex-concave and can have multiple 323 local equilibria or stationary points [38]. The adversary's objective also does not have a strong 324 dual that can be solved through conic programs—a standard practice in DRO literature [42]. Thus, 325 we provide an algorithm where both learner and adversary optimize BR-DRO iteratively through 326 stochastic gradient ascent/descent (Algorithm 1). The adversary's action space $\mathcal{W}(\gamma)$ is constrained 327 either with an information bottleneck penalty by setting $\beta_{\rm vib} \neq 0$ or l_2 norm penalty by setting 328 $\beta_{l_2} \neq 0$ in equation 5 below. While we can choose to constrain the adversary with both forms of 329 constraints simultaneously we find that in practice picking only one of them for a given problem 330 instance helps with tuning the degree of constraint. For more details on the architecture and other 331 details see Appendix E. 332

$$\min_{\boldsymbol{\theta}_h \in \Theta_h} \langle l(\boldsymbol{\theta}_h), \boldsymbol{\theta}_w^* \rangle_P \quad \text{s.t.} \quad \boldsymbol{\theta}_w^* = \operatorname*{arg\,max}_{\boldsymbol{\theta}_w \in \Theta_w} \quad L_{adv}(\boldsymbol{\theta}_w; \boldsymbol{\theta}_h, \beta_{\text{vib}}, \beta_{l_2}, \eta) \tag{5}$$

$$L_{\text{adv}}(\boldsymbol{\theta}_{w};\boldsymbol{\theta}_{h},\beta_{\text{vib}},\beta_{l_{2}},\eta) = \langle l(\boldsymbol{\theta}_{h}) - \eta,\boldsymbol{\theta}_{w} \rangle_{P} - \beta_{\text{vib}} \mathbb{E}_{P}\text{KL}(p(\mathbf{z} \mid \mathbf{x};\boldsymbol{\theta}_{w}) \mid\mid \mathcal{N}(\mathbf{0},\mathbf{I_{d}})) - \beta_{\mathbf{l_{2}}} \|\boldsymbol{\theta}_{\mathbf{w}}\|_{\mathbf{2}}^{2}$$

Training. For each example, the adversary takes as input: (i) the last layer output of the current 333 learner's feature network; and (ii) the input label. The adversary then outputs a weight (in [0, 1]). The 334 idea of applying the adversary directly on the learner's features (instead of the original input) is based 335 on recent literature [45, 28] that suggests re-training the prediction head is sufficient for robustness to 336 shifts. The adversary tries to maximize weights on examples with value $\geq \eta$ (hyperparameter) and 337 minimize on others. For the learner, in addition to the example it takes as input the adversary assigned 338 weight for that example from the previous round and uses it to reweigh its loss in a minibatch. Both 339 players are updated in a round (Algorithm 1). 340

²We use θ_h, θ_w and $l(\theta_h)$ to denote $w(\theta_w; (\mathbf{x}, \mathbf{y})), h(\theta_h; \mathbf{x})$ and $l(h(\theta_h; \mathbf{x}), \mathbf{y})$ respectively.

341 **B** Theoretical Analysis

The main objective of our analysis of BR-DR0 is to show how adding a bitrate constraint on the adversary can: (i) give us tighter statistical estimates of the worst risk; and (ii) control the pessimism (excess risk) of the learned solution. First, we provide worst risk generalization guarantees using the PAC-Bayes framework [15], along with a result for kernel adversary. Then, we provide convergence rates and pessimism guarantees for the solution found by our online solver for a specific instance of $W(\gamma)$. For both these, we analyze the constrained form of the conditional value at risk (CVaR) DRO objective [31] below.

Bitrate-Constrained CVaR DRO. When the uncertainty set Q is defined by the set of all distributions Q that have bounded likelihood *i.e.*, $||q/p||_{\infty} \leq 1/\alpha_0$, we recover the original CVaR DRO objective [19]. The bitrate-constrained version of CVaR DRO is given in equation 6 (see Appendix G for derivation). Note that, slightly different from Section 2, we define W as the set of all measurable functions $w: \mathcal{X} \times \mathcal{Y} \mapsto [0, 1]$, since the other convex restrictions in equation 1 are handled by dual variable η . As in Section 2, $W(\gamma)$ is derived from W using Definition 2.4. In equation 6, if we replace the bitrate-constrained class $W(\gamma)$ with the unrestricted W then we recover the variational form of unconstrained CVaR DRO in Duchi et al. [17].

$$\mathcal{L}_{\text{evar}}^*(\gamma) = \inf_{h \in \mathcal{H}, \eta \in \mathbb{R}} \sup_{w \in \mathcal{W}(\gamma)} R(h, \eta, w) \text{ where, } R(h, \eta, w) = (1/\alpha_0) \langle l(h) - \eta, w \rangle_P + \eta$$
 (6)

357 B.1 Worst risk estimation bounds for BR-DRO.

Since we are only given a finite sampled dataset $\mathcal{D} \sim P^n$, we solve the objective in equation 6 using the empirical distribution \hat{P}_n . We denote the plug-in estimates as \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$. This incurs an estimation error for the true worst risk. But when we restrict our adversary to $\Delta(\mathcal{W}, \gamma)$, for a fixed learner h we reduce the worst-case risk estimation error which scales with the bitrate KL($\cdot || \pi$) of the solution (deviation from prior π). Expanding this argument to every learner in \mathcal{H} , with high probability we also reduce the estimation error for the worst risk of \hat{h}_D^{γ} . Theorem B.1 states this generalization guarantee more precisely.

Theorem B.1 (worst-case risk generalization). With probability $\geq 1 - \delta$ over $\mathcal{D} \sim P^n$, the worst

bitrate-constrained α_0 -CVaR risk for \hat{h}_D^{γ} can be upper bounded by the following oracle inequality:

$$\sup_{w \in \mathcal{W}(\gamma)} R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w) \lesssim \mathcal{L}_{cvar}^*(\gamma) + \frac{M}{\alpha_0} \sqrt{\left(\gamma + \log\left(\frac{1}{\delta}\right) + (d+1)\log\left(\frac{L^2n}{\gamma}\right) + \log n\right)} / (2n-1),$$

when $l(\cdot, \cdot)$ is [0, M]-bounded, L-Lipschitz and \mathcal{H} is parameterized by convex set $\Theta \subset \mathbb{R}^d$.

Informally, Theorem B.1 tells us that bitrate-constraint γ gracefully controls the estimation error 368 $\mathcal{O}(\sqrt{(\gamma + \mathcal{C}(\mathcal{H}))/n})$ (where $\mathcal{C}(\mathcal{H})$ is a complexity measure) if we know that Assumption 2.5 is 369 satisfied. While this only tells us that our estimator is consistent with $\mathcal{O}_p(1/\sqrt{n})$, the estimate may 370 itself be converging to a degenerate predictor, *i.e.*, $\mathcal{L}^*_{cvar}(\gamma)$ may be very high. For example, if the 371 adversary can cleanly separate mislabeled points even after the bitrate constraint, then presumably 372 these noisy points with high losses would be the ones mainly contributing to the worst risk, and 373 up-weighting these points would result in a learner that has memorized noise. Thus, it becomes 374 equally important for us to analyze the excess risk (or the pessimism) for the learned solution. Since 375 this is hard to study for any arbitrary bitrate-constrained class $\mathcal{W}(\gamma)$, we shall do so for the specific 376 class of reproducing kernel Hilbert space (RKHS) functions. 377

Special case of bounded RKHS. Let us assume there exists a prior Π such that $\mathcal{W}(\gamma)$ in Definition 2.4 is given by an RKHS induced by Mercer kernel $k : \mathcal{X} \times \mathcal{X} \mapsto \mathbb{R}$, s.t. the eigenvalues of the kernel operator decay polynomially, *i.e.*, $\mu_j \leq j^{-2/\gamma}$ ($\gamma < 2$). Then, if we solve for \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$ by doing kernel ridge regression over norm bounded ($||f||_{\mathcal{W}(\gamma)} \leq B \leq 1$) smooth functions f then we can control: (i) the pessimism of the learned solution; and (ii) the generalization error (Theorem B.2). Formally, we refer to pessimism for estimates \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$ as excess risk defined as:

excess risk :=
$$\sup_{w \in \mathcal{W}(\gamma)} |\inf_{h,\eta} R(h,\eta,w) - R(\hat{h}_D^{\gamma},\hat{\eta}_D^{\gamma},w)|.$$
(7)

Theorem B.2 (bounded RKHS). For l, H in Theorem B.1, and for $W(\gamma)$ described above $\exists \gamma_0 s.t.$ 384 for all sufficiently bitrate-constrained $W(\gamma)$ i.e., $\gamma \leq \gamma_0$, w.h.p. $1 - \delta$ worst risk generalization error 385

is $\mathcal{O}\left((1/n)\left(\log(1/\delta) + (d+1)\log(nB^{-\gamma}L^{\gamma/2})\right)\right)$ and the excess risk is $\mathcal{O}(B)$ for $\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}$ above. 386

Thus, in the setting described above we have shown how bitrate-constraints given indirectly by γ , R 387 can control both the pessimism and statistical estimation errors. Here, we directly analyzed the 388 estimates \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$ but did not describe the specific algorithm used to solve the objective in equation 6 with \hat{P}_n . Now, we look at an iterative online algorithm to solve the same objective and see how 389 390 bitrate-constraints can also influence convergence rates in this setting. 391

Convergence and excess risk analysis for an online solver. **B.2** 392

In the following, we provide an algorithm to solve the objective in equation 6 and analyze how 393 bitrate-constraint impacts the solver and the solution. For convex losses, the min-max objective in 394 equation 6 has a unique solution and this matches the unique Nash equilibrium for the generic online 395 algorithm (game) we describe (Lemma B.3). The algorithm is as follows: Consider a two-player 396 zero-sum game where the learner uses a no-regret strategy to first play $h \in \mathcal{H}, \eta \in \mathbb{R}$ to minimize 397 $\mathbb{E}_{w \sim \delta} R(h, \eta, w)$. Then, the adversary plays follow the regularized leader (FTRL) strategy to pick 398 distribution $\delta \in \Delta(\mathcal{W}(\gamma))$ to maximize the same. Our goal is to analyze the bitrate-constraint γ 's 399 effect on the above algorithm's convergence rate and the pessimistic nature of the solution found. 400 For this, we need to first characterize the bitrate-constraint class $\mathcal{W}(\gamma)$. If we assume there exists a 401 prior Π such that $\mathcal{W}(\gamma)$ is Vapnik-Chervenokis (VC) class of dimension $O(\gamma)$, then in Theorem B.4, 402 we see that the iterates of our algorithm converge to the equilibrium (solution) in $\mathcal{O}(\sqrt{\gamma \log n/T})$ 403 steps. Clearly, the degree of bitrate constraint can significantly impact the convergence rate for a 404 generic solver that solves the constrained DRO objective. Theorem B.4 also bounds the excess risk 405 (equation 7) on \hat{P}_n . 406

Lemma B.3 (Nash equilibrium). For convex l(h), $l(h) \in [0, M]$, the objective in equation 6 has a 407 unique solution which is also the Nash equilibrium of the game above when played over compact sets 408 $\mathcal{H} \times [0, M], \Delta(\mathcal{W}, \gamma)$. We denote this equilibrium as $h_D^*(\gamma), \eta_D^*(\gamma), \delta_D^*(\gamma)$. 409

Theorem B.4. At time step t, if the learner plays (h_t, η_t) with no-regret and the adversary plays δ_t with FTRL strategy that uses a negative entropy regularizer on δ then average iterates $(\bar{h}_T, \bar{\eta}_T, \bar{\delta}_T) =$ 410 411 $(1/T) \sum_{t=1}^{T} (h_t, \eta_t, \delta_t)$ converge to the equilibrium $(h_D^*(\gamma), \eta_D^*(\gamma), \delta_D^*(\gamma))$ at rate $\mathcal{O}(\sqrt{\gamma \log n/T})$. Further the excess risk defined above is $\mathcal{O}((M/\alpha_0)(1-\frac{1}{n^{\gamma}}))$. 412

413

С **Detailed experiments** 414

Our experiments aim to evaluate the performance of BR-DRO and compare it with ERM and group 415 shift robustness methods that do not require group annotations for training examples. We conduct 416 empirical analyses along the following axes: (i) worst group performance on datasets that exhibit 417 known spurious correlations; (ii) robustness to random label noise in the training data; (iii) aver-418 age performance on hybrid covariate shift datasets with unspecified groups; and (iv) accuracy in 419 identifying minority groups. See Appendix F for additional experiments and details. 420

421 **Baselines.** Since our objective is to be robust to group shifts without group annotations on training examples, we explore baselines that either optimize for the worst minority group (CVaR DRO [31]) 422 or use training losses to identify specific minority points (LfF [41], JTT [35]). Group DRO [46] is 423 treated as an oracle. We also compare with the simple re-weighting baseline (RWY) proposed by 424 Idrissi et al. [26]. 425

Implementation details. We train using Resnet-50 [24] for all methods and datasets except Civil-426 Comments, where we use BERT [58]. For our VIB adversary, we use a 1-hidden layer neural network 427 encoder and decoder (one for each label). As mentioned in Section 2, the adversary takes as input 428 the learner model's features and the true label to generate weights. All implementation and design 429 choices for baselines were adopted directly from Liu et al. [35], Idrissi et al. [26]. We provide model 430 selection methodology and other details in Appendix F. 431

Datasets. For experiments in the known groups and label noise settings we use: (i) Waterbirds [54] 432 (background is spurious), CelebA [36] (binary gender is spuriously correlated with label "blond"); and 433

Method	FM	oW	Camelyon17	
	Avg	W-Reg	Avg	
ERM	53.3 (0.1)	32.4 (0.3)	70.6 (1.6)	
JTT [35]	52.1 (0.1)	31.8 (0.2)	66.3 (1.3)	
LfF [41]	49.6 (0.2)	31.0 (0.3)	65.8 (1.2)	
RWY [26]	50.8 (0.1)	30.9 (0.2)	69.9 (1.3)	
Group DRO [46]	51.9 (0.2)	30.4 (0.3)	68.5 (0.9)	
CVaR DRO [31]	51.5 (0.1)	31.0 (0.3)	66.8 (1.3)	
BR-DRO (VIB) (ours)	52.0 (0.2)	31.8 (0.2)	70.4 (1.5)	
BR-DRU (l_2) (ours)	55.1 (0.1)	52.5 (0.2)	/1.2 (1.0)	

Table 2: Average (Avg) and worst region (W-Reg for FMoW) test accuracies on Camelyon17 and FMoW.

CivilComments (WILDS) [11] where the task is to predict "toxic" texts and there are 16 predefined groups [29]. We use FMoW and Camelyon17 [29] to test methods on datasets that do not have explicit group shifts. In FMoW the task is to predict land use from satellite images where the training/test set comprises of data before/after 2013. Test involves both subpopulation shifts over regions (*e.g.*, Africa, Asia) and domain generalization over time (year). Camelyon17 presents a domain generalization problem where the task is to detect tumor in tissue slides from different sets of hospitals in train and test sets.

441 C.1 More experiments on robustness to noise.

To further verify our claims, we set up a noisily labeled synthetic dataset (see Appendix F for details). 442 In Figure 2a we plot training samples as well as the solutions learned by BR-DRO and and JTT on 443 synthetic data. In Figure 1(right) we also plot exactly which points are upweighted by BR-DRO and 444 JTT. Using both figures, we note that JTT mainly upweights the noisy points (in red) and memorizes 445 them using \mathbf{x}_{noise} . Without any weights on minority, it memorizes them as well and learns component 446 along spurious feature. On the contrary, when we restrict the adversary with BR-DRO to be sparse 447 $(l_1 \text{ penalty})$, it only upweights minority samples, since no sparse predictor can separate noisy points 448 in the data. Thus, the learner can no longer memorize the upweighted minority and we recover the 449 robust predictor along core feature. 450

451 C.2 How does BR-DRO perform on more general covariate shifts?

In Figure 2 we report the average test accuracies for BR-DRO and baselines on the hybrid dataset 452 FMoW and domain generalization dataset Camelyon 17. In (\cdot) we report the standard error of the 453 mean accuracy across five runs. Given its hybrid nature, on FMoW we also report worst region 454 accuracy. First, we note that on these datasets group shift robustness baselines do not do better than 455 ERM. Some are either too pessimistic (e.g., CVaR DRO), or require heavy assumptions (e.g., Group 456 DRO) to be robust to domain generalization. This is also noted by Gulrajani and Lopez-Paz [22]. 457 Next, we see that BR-DRO (l_2 version) does better than other group shift baselines on both both worst 458 region and average datasets and matches ERM performance on Camelyon17. One explanation could 459 be that even though these datasets test models on new domains, there maybe some latent groups 460 defining these domains that are simple and form a part of latent subpopulation shift. Investigating 461 this claim further is a promising line of future work. 462

463 C.3 What fraction of minority is recovered by Bitrate-Constrained DRO?

We claim that our less pessimistic objective can more accurately recover (upsample) the true minority 464 group if indeed the minority group is simple (see Assumption 2.5 for our definition of simple). In 465 this section, we aim to verify this claim. If we treat examples in the top 10% (chosen for post hoc 466 analysis) fraction of examples as our predicted minorities, we can check precision and recall of this 467 decision on the Waterbirds and CelebA datasets. Figure 3 plots these metrics at each training epoch 468 for BR-DRO (with varying β_{vib}), JTT and CVaR DRO. Precision of the random baseline tells us the 469 true fraction of minority examples in the data. First we note that BR-DRO consistently performs much 470 better on this metric than unconstrained CVaR DRO. In fact, as we reduce strength of β_{vib} we recover 471 precision/recall close to the latter. This controlled experiment shows that the bitrate constraint is 472



Figure 3: By considering the fraction of points upweighted by our adversary (top 10%) as the positive class we analyze the precision and recall of this class with respect to the minority group. and do the same for JTT, random baseline and CVaR DRO. BR-DRO achieves highest precision and matches recall with JTT asymptotically. We also find that increasing bitrate constraint β_{vib} helps improving precision/recall.

helpful (and very much needed) in practice to identify rare simple groups. In Figure 3 we observe that
asymptotically, the precision of BR-DRO is better than JTT on both datasets, while the recall is similar.
Since importance weighting has little impact in later stages with exponential tail losses [52, 13], other
losses (*e.g.*, polytail Wang et al. [56]) may further improve the performance of BR-DRO as it gets
better at identifying the minority classes when trained longer.

478 **D** Related Work

Prior works in robust ML [e.g., 32, 33, 20] address various forms of adversarial or structured shifts.
We specifically review prior work on robustness to group shifts. While those based on DRO optimize
for worst-case shifts in an explicit uncertainty set, the robust set is implicit for some others, with most
using some form of importance weighting.

Distributionally robust optimization (DRO). DRO methods generally optimize for worst-case 483 performance on joint (x, y) distributions that lie in an f-divergence ball (uncertainty set) around the 484 training distribution [7, 44, 8, 9, 40, 17, 19]. Hu et al. [25] highlights that the conservative nature 485 of DRO may lead to degenerate solutions when the unrestricted adversary uniformly upweights all 486 misclassified points. Sagawa et al. [46] proposes to address this by limiting the adversary to shifts that 487 only differ in marginals over predefined groups. However, in addition to it being difficult to obtain 488 this information, Kearns et al. [27] raise "gerrymandering" concerns with notions of robustness that 489 fix a small number of groups apriori. While they propose a solution that looks at exponentially many 490 subgroups defined over protected attributes, our method does not assume access to such attributes 491 and aims to be fair on them as long as they are realized by simple functions. Finally, Zhai et al. 492 [60] avoid conservative solutions by solving the DRO objective over randomized predictors learned 493 through boosting. We consider deterministic and over-parameterized learners and instead constrain 494 the adversary's class. 495

Constraining the DRO uncertainty set. In the marginal DRO setting, Duchi et al. [18] limit 496 the adversary via easier-to-control reproducing kernel hilbert spaces (RKHS) or bounded Hölder 497 continuous functions [34, 57]. While this reduces the statistical error in worst risk estimation, the 498 size of the uncertainty set (scales with the data) remains too large to avoid cases where an adversary 499 can re-weight mislabeled and hard examples from the majority set [14]. In contrast, we restrict the 500 adversary even for large datasets where the estimation error would be low, as this would reduce 501 excess risk when we only care about robustness to rare sub-populations defined by simple functions. 502 Additionally, while their analysis and method prefers the adversary's objective to have a strong dual, 503 we show empirical results on real-world datasets and generalization bounds where the adversary's 504 objective is not necessarily convex. 505

Robustness to group shifts without demographics. Recent works [50, 16, 5] that aim to achieve group robustness without access to group labels employ various heuristics where the robust set is implicit while others require data from multiple domains [3, 59] or ability to query test samples [30]. Liu et al. [35] use training losses for a heavily regularized model trained with empirical risk minimization (ERM) to directly identify minority data points with higher losses and re-train on the dataset that up-weights the identified set. Nam et al. [41] take a similar approach. Other methods [26] propose simple baselines that subsample the majority class in the absence of group demographics and the majority group in its presence. Hashimoto et al. [23] find DRO over a χ^2 -divergence ball can reduce the otherwise increasing disparity of per-group risks in a dynamical system. Since it does not use features to upweight points (like BR-DRO) it is vulnerable to label noise. Same can be said about some other works (*e.g.*, [35, 41]).

Importance weighting in deep learning. Finally, numerous works [17, 31, 33, 43] enforce robustness by re-weighting losses on individual data points. Recent investigations [52, 13, 37] reveal that such objectives have little impact on the learned solution in interpolation regimes. One way to avoid this pitfall is to train with heavily regularized models [46, 47] and employ early stopping. Another way is to subsample certain points, as opposed to up-weighting [26]. In this work, we use both techniques while training our objective and the baselines, ensuring that the regularized class is robust to shifts under misspecification [57].

524 E BR-DRO algorithm

If the bitrate constraint is applied via the KL term in equation 5, we implement the adversary as a variational information bottleneck [2] (VIB), where the KL divergence with respect to a standard Gaussian prior controls the bitrate of the adversary's feature set $\mathbf{z} \sim p(\mathbf{z} \mid \mathbf{x}; \boldsymbol{\theta}_w)$. Increasing β_{vib} can be seen as enforcing lower bitrate features *i.e.*, reducing γ in $\mathcal{W}(\gamma)$ (smaller value of KL($\delta \mid\mid \pi$) in the primal formulation in Definition 2.4). If the constraint is applied via the l_2 term we implement the adversary as a linear layer. In some cases (*e.g.*, Section 3) we use a sparsity constraint (l_1 norm) on the linear adversary.

Algorithm 1: Bitrate-Constraint DRO (Online Algorithm)Input: Adversary VIB penalty β_{vib} ; Step sizes η_l, η_w ; Dataset $\mathcal{D} = (\mathbf{x}_i, \mathbf{y}_i)_{i=1}^n$ Initialize $\theta_h^{(1)}$ and $\theta_w^{(1)}$ for t = 1, ..., T doFrom \mathcal{D} , sample $\mathbf{x}, \mathbf{y} \sim \mathcal{D}$ $\theta_h^{(t+1)} \leftarrow \Pi_{\Theta_h} \left(\theta_h^{(t)} - \eta_h \nabla_{\theta_h} \left[l(\theta_h^{(t)}(\mathbf{x}), \mathbf{y}) \cdot \theta_w(\mathbf{x}, \mathbf{y}) \right] \right)$ /* Update $\theta_h */$ $\theta_w^{(t+1)} \leftarrow \Pi_{\Theta_w} \left(\theta_w^{(t)} + \eta_w \nabla_{\theta_w} L_{adv}(\theta_w^{(t)}; \theta_h^{(t)}, \beta_{vib}, \beta_{l_2}, \eta) \right)$ endOutput: $\bar{\theta}_h = \frac{1}{T} \sum_{t=1}^T \theta_h^{(t)}, \bar{\theta}_w = \frac{1}{T} \sum_{t=1}^T \theta_w^{(t)}$

⁵³² F Additional empirical results and other experiment details

533 F.1 Hyper-parameter tuning methodology

There are two ways in which we tune hyperparameters on datasets with known groups (CelebA, 534 Waterbirds, CivilComments): (i) on average validation performance; (ii) worst group accuracy. The 535 former does not use group annotations while the latter does. Similar to prior works [35, 26] we note 536 that using group annotations (on a small validation set) does improve performance. In Table 3 we 537 report our study which varies the the fraction p of group labels that are available at test time. For 538 each setting of p, we do model selection by taking weighted (by p) mean over two entities (i) average 539 validation on all samples, (ii) worst group validation on a fraction p of minority samples. In the case 540 where p = 0, we only use average validation. We report our results on CelebA and Waterbirds dataset. 541 For the two WILDS datasets we tune hyper-parameters on OOD Validation set. 542

543 F.2 Synthetic dataset details

We follow the explicit-memorization setup in Sagawa et al. [47] which we summarize here briefly. Let input $\mathbf{x} = [\mathbf{x}_{core}, \mathbf{x}_{spu}, \mathbf{x}_{noise}]$ where $\mathbf{x}_{core} \mid y \sim \mathcal{N}(y, \sigma_{core}^2), \mathbf{x}_{spu} \mid a \sim \mathcal{N}(a, \sigma_{spu}^2)$ and $\mathbf{x}_{noise} \sim \mathcal{N}(0, (\sigma_{noise}^2 \mathbf{I_d})/\mathbf{d})$. Here $a \in \{-1, 1\}$ refers to a spurious attribute, and label is $y \in \{-1, 1\}$, We set a = y with probability $P(_{maj}) = 1 - P(_{min})$. The level of correlation between a and y is controlled by $P(_{maj})$. Additionally, we flip true label with probability $P(_{noise})$.

	Waterbirds				CelebA			
Method	p = 0.0	p=0.02	p = 0.05	p = 0.1	p = 0.0	p = 0.02	p = 0.05	p = 0.1
JTT	62.7	73.9	77.3	84.4	42.1	68.3	80.5	80.3
CVaR DRO	63.9	65.8	72.6	74.1	33.6	40.4	60.4	63.2
LfF	48.6	58.9	70.3	79.5	34.0	58.9	60.0	78.3
BR-DRO (VIB)	69.3	77.6	76.1	84.9	52.4	71.2	80.3	79.9
BR-DRO (l_2)	68.9	75.2	79.4	86.1	55.8	63.5	74.6	80.4

Table 3: We check to what extent fraction of group annotations in the training data affect performance. For each dataset and method, we tune its hyper-parameters on the average validation and worst group (only on the small fraction p that is available). We see that while all methods consistently improve as we increase group annotations and tune for worst group accuracy on the annotated samples, BR-DRO does do better that prior works when tuned on just average validation (p = 0.). At the same time, we note that this still does not match the performance of BR-DRO when tuned on worst group validation (seen in Table 1).

549 F.3 Degree of constraint

In Figure 4 we see how worst group performance varies on Waterbirds and CelebA as a function of increasing constraint. We also plot average performance on the Camelyon dataset. We mainly note that for either of the constraint implementations, only when we significantly increase the capacity do we actually see the performance of BR-DR0 improve. The effect is more prominent on groups shift datasets with simple groups (Waterbirds, CelebA). Under less restrictive capacity constraints we note that its performance is similar to CVaR DRO (see Figure 3). This is expected since CVaR DRO is the completely unconstrained version of our objective.



Figure 4: **Optimal bitrate-constraints for robustness to distributions shifts:** For two versions of capacity control: KL, l_2 penalty (see Section 2) we show how worst group performance on Waterbirds, CelebA and average performance on Camelyon test sets improves with increasing constraints under either VIB (β_{vib}) or linear (β_{l_2}) adversaries.

557 F.4 Hyper-parameter details.

For all hyper-parameters of prior methods we use the ones state in their respective prior works. The 558 implementation Group DRO, JTT, CVaR DRO is borrowed from the implementation made public 559 560 by authors of Liu et al. [35]. For datasets Waterbirds, CelebA and CivilComments we choose the hyper-parameters (whenever applicable) learning rate, batch size, weight decay on learner, optimizer, 561 early stopping criterion, learning rate schedules used by Liu et al. [35] for their implementation 562 of CVaR DRO method. For datasets FMoW and Camelyon17 we choose values for these hyper-563 parameters to be the ones used by Koh et al. [29] for the ERM baseline. Details on BR-DRO specific 564 hyper-parameters that we tuned are in Table 4. Also, note that we release our implementation with 565 this submission. 566

567 G Omitted Proofs

First we shall state some a couple of technical lemmas that we shall refer to at multiple points. Then, we prove our theoretical claims in our analysis in Appendix B, in the order in which they

Hyper-parameter	Waterbirds	CelebA	CivilComments	FMoW	Camelyon17
learning rate for adversary	0.01	0.05	0.001	0.02	0.01
threshold η	0.05	0.05	0.1	0.1	0.1
$\beta_{\rm vib}$	0.1	0.1	0.02	0.005	0.005
β_{l_2}	0.01	0.005	0.005	0.02	0.005

Table 4: Hyperparameters for our method on different datasets (tuned on worst group validation performance). Note, that the threshold η here is the top x% fraction.

appear. Before we get into those we provide proof for our Corollary 2.2 and the derivation of Bitrate-Constrained CVaR DRO in Equation 6.

Lemma G.1 (Hoeffding bound [55]). Let X_1, \ldots, X_n be a set of μ_i centered independent sub-Gaussians, each with parameter σ_i . Then for all $t \ge 0$, we have

$$\mathbb{P}\left[\frac{1}{n}\sum_{i=1}^{n}(X_i-\mu_i)\geq t\right]\leq \exp\left(-\frac{n^2t^2}{2\sum_{i=1}^{n}\sigma_i^2}\right).$$
(8)

Lemma G.2 (Lipschitz functions of Gaussians [55]). Let X_1, \ldots, X_n be a vector of iid Gaussian variables and $f : \mathbb{R}^n \to \mathbb{R}$ be L-Lipschitz with respect to the Euclidean norm. Then the random variable $f(X) - \mathbb{E}[f(X)]$ is sub-Gaussian with parameter at most L, thus:

$$\mathbb{P}[|f(X) - \mathbb{E}[f(X)]| \ge t] \le 2 \cdot \exp\left(-\frac{t^2}{2L^2}\right), \quad \forall t \ge 0.$$
(9)

577 G.1 Proof of Corollary 2.2

Let us recall the definition of a well defined group structure. For a pair of measures $Q \ll P$ we say $\mathcal{G}(P,Q)$ is well defined if given there exists a set of disjoint measurable sets $\mathcal{G}_{P,Q} = \{G_k\}_{k=1}^K$ such that $G_k \in \Sigma$, $Q(G_k) > 0$, $Q(\mathcal{G}(P,Q)) = 1$ and we have:

$$K = \min\{|\{G_1, \dots, G_M\}|: \ p(\mathbf{x}, \mathbf{y} \mid G_m) = q(\mathbf{x}, \mathbf{y} \mid G_m) > 0, \forall (\mathbf{x}, \mathbf{y}) \in G_m \ \forall m \in [M]\}$$

Now by definition K is finite. Thus if there exists two well defined group structures $\mathcal{G}_1(P,Q)$ and $\mathcal{G}_2(P,Q)$ for the same pair P, Q then it must be the case that $K = \mathcal{G}_1(P,Q) = \mathcal{G}_2(P,Q)$.

Then, there must exist $G \in \mathcal{G}_1(P,Q)$ such that Q(G) > 0 and $G', G'' \in \mathcal{G}_2(P,Q)$ where Q(G'), Q(G'') > 0 and $Q(G \cap G'), Q(G \cap G'') > 0$.

Note that since $G, G', G'' \in \Sigma$ that is closed under countable unions, we have that $G \cap G'$ and $G \cap G''$ are two sets where $q(\mathbf{x}, \mathbf{y}) > 0 \ \forall (\mathbf{x}, \mathbf{y}) \in G \cap G', G \cap G''$.

Let $(\mathbf{x}_1, \mathbf{y}_1) \in (G \cap G')$ and $(\mathbf{x}_2, \mathbf{y}_2) \in (G \cap G'')$. From definition we know that $q(\mathbf{x}_2, \mathbf{y}_2), q(\mathbf{x}_1, \mathbf{y}_1) > 0$ and . Since both $(\mathbf{x}_1, \mathbf{y}_1)$ and $(\mathbf{x}_2, \mathbf{y}_2)$ are in G we have that:

$$q(\mathbf{x}_1, \mathbf{y}_1) = \frac{Q(G)}{P(G)} \cdot p(\mathbf{x}_1, \mathbf{y}_1) = \frac{Q(G')}{P(G')} \cdot p(\mathbf{x}_1, \mathbf{y}_1)$$
(10)

$$q(\mathbf{x}_2, \mathbf{y}_2) = \frac{Q(G)}{P(G)} \cdot p(\mathbf{x}_2, \mathbf{y}_2) = \frac{Q(G'')}{P(G'')} \cdot p(\mathbf{x}_2, \mathbf{y}_2)$$
(11)

Thus, we can conclude that $\frac{Q(G')}{P(G')} = \frac{Q(G'')}{P(G'')}$. This implies that $G' \cup G''$ also satisfies the following that $Q(G' \cup G'') > 0$ and $q(\mathbf{x}, \mathbf{y} \mid G' \cup G'') = p(\mathbf{x}, \mathbf{y} \mid G' \cup G'')$.

Thus, we can construct a new $\mathcal{G}_3(P,Q) = \{G \in \mathcal{G}_2(P,Q) : G \notin \{G',G''\}\} \cup \{G' \cup G''\}$. Clearly, $\mathcal{G}_3(P,Q)$ satisfies all group structure properties and is smaller than $\mathcal{G}_2(P,Q)$. Thus, we arrive at a

 $\mathcal{G}_3(P,Q)$ satisfies all group structure properties and is smaller than $\mathcal{G}_2(P,Q)$. Thus, we ar contradiction which proves the claim that $\mathcal{G}(P,Q)$ is indeed unique whenever well defined.

$\frac{1}{2}$ contradiction which proves the crain that g(r, q) is indeed unique whenever we defined

594 G.2 Derivation of Bitrate-Constrained CVaR DRO in equation 6

Recall that we define \mathcal{W} as the set of all measurable functions $w : \mathcal{X} \times \mathcal{Y} \mapsto [0, 1]$, since the other convex restrictions in equation 1 are handled by dual variable η . As in Section 2, $\mathcal{W}(\gamma)$ is derived from the new \mathcal{W} using Definition 2.4. With that let us first state the CVaR objective [31].

$$\mathcal{L}_{\text{cvar}}(h, P) \coloneqq \sup_{q} \int_{\mathcal{X} \times \mathcal{Y}} q(\mathbf{x}, \mathbf{y}) \cdot l(h)$$

s.t. $q \ge 0, \ \|q/p\|_{\infty} \le (1/\alpha_0), \ \int_{\mathcal{X} \times \mathcal{Y}} q(\mathbf{x}, \mathbf{y}) = 1$ (12)

The objective in q is linear with convex constraints, and has a strong dual (see Duchi et al. [17], Boyd et al. [12] for the derivation) which is given by:

$$\inf_{\eta \in \mathbb{R}} \left\{ \frac{1}{\alpha_0} \mathbb{E}_P(l(h) - \eta)_+ + \eta \right\} \\
= \inf_{\eta \in \mathbb{R}} \left\{ \frac{1}{\alpha_0} \langle (l(h) - \eta)_+, \mathbb{1} \rangle_P + \eta \right\} \\
= \inf_{\eta \in \mathbb{R}} \left\{ \frac{1}{\alpha_0} \langle (l(h) - \eta), \mathbb{1}(l(h) - \eta \ge 0) \rangle_P + \eta \right\}$$
(13)

$$= \inf_{\eta \in \mathbb{R}} \sup_{w \in \mathcal{W}} \left\{ \frac{1}{\alpha_0} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
(14)

The last equality is true since the set $\mathbb{1}(l(h) - \eta \ge 0)$ is measurable under P (based on our setup in Appendix B). Note that for any h, the objective $\frac{1}{\alpha_0} \langle (l(h) - \eta), w \rangle_P + \eta$ is linear in w, and η . If we further assume the loss l(h) to be the l_{0-1} loss, it is bounded, and thus the optimization over η can be restricted to a compact set. Next, W is also a compact set of functions since we restrict our solvers to measurable functions that take values bounded in [0, 1].

$$\mathcal{L}_{\text{cvar}}(h, P) = \inf_{\eta \in \mathbb{R}} \sup_{w \in \mathcal{W}} \left\{ \frac{1}{\alpha_0} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
(15)

The above objective is precisely the Bitrate-Constrained CVaR DRO objective we have in equation 6. Later in the Appendix we shall need an equivalent form of the objective which we shall derive below.

⁶⁰⁷ We can now invoke the Weierstrass' theorem in Boyd et al. [12] to give us the following:

$$\mathcal{L}_{\text{cvar}}(h, P) = \inf_{\eta \in \mathbb{R}} \sup_{w \in \mathcal{W}} \left\{ \frac{1}{\alpha_0} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
$$= \frac{1}{\alpha_0} \sup_{w \in \mathcal{W}} \left\{ \inf_{\eta \in \mathbb{R}} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
(16)

Now, the final objective $\inf_{h \in \mathcal{H}} \mathcal{L}_{cvar}(h, P)$ is given by:

$$\frac{1}{\alpha_0} \inf_{h \in \mathcal{H}} \sup_{w \in \mathcal{W}} \left\{ \inf_{\eta \in \mathbb{R}} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
(17)

In the above equation we can now replace the unconstrained class \mathcal{W} with our bitrate-constrained class $\mathcal{W}(\gamma)$ to get the following:

$$\frac{1}{\alpha_0} \inf_{h \in \mathcal{H}} \sup_{w \in \mathcal{W}(\gamma)} \left\{ \inf_{\eta \in \mathbb{R}} \langle (l(h) - \eta), w \rangle_P + \eta \right\}$$
(18)

611 G.3 Proof of Theorem B.1

- ⁶¹² For convenience we shall first restate the Theorem here.
- Theorem G.3 ([restated). worst-case risk generalization] With probability $\geq 1 \delta$ over sample
- 614 $\mathcal{D} \sim P^n$, the worst risk for \hat{h}_D^{γ} can be upper bounded by the following oracle inequality:

$$\sup_{w \in \Delta(\mathcal{W},\gamma)} R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w) - \mathcal{L}_{cvar}^*(\gamma) \lesssim \frac{M}{\alpha_0} \sqrt{\left(\gamma + \log\left(\frac{1}{\delta}\right) + (d+1)\log\left(\frac{L^2n}{\gamma}\right) + \log n\right)} / (2n-1)$$

- 615 when $l(\cdot, \cdot)$ is [0, M]-bounded, L-Lipschitz and \mathcal{H} is parameterized by convex set $\Theta \subset \mathbb{R}^d$.
- 616 The overview of the proof can be split into two parts:
- For each learner, first obtain the oracle PAC-Bayes [39] worst risk generalization guarantee over the adversary's action space $\Delta(W, \gamma)$.
- Then, apply uniform convergence bounds using a union bound over a covering of the class \mathcal{H} to get the final result.

Intuition: The only tricky part lies in the fact that oracle PAC-Bayes inequality would not give us arbitrary control over the generalization error for each learner, which we would typically get in Hoeffding type bounds. Hence, we need to ensure that the the worst risk generalization rate decays faster than how the size of the covering would increase for a ball of radius defined by the worst generalization error.

Now, we shall invoke the following PAC-Bayes generalization guarantee stated (Lemma G.4) since $R(h, \eta, w) \in [0, M/\alpha_0]$.

Lemma G.4 (PAC-Bayes [15, 39]). With probability $\geq 1 - \delta$ over choice of dataset \mathcal{D} of size n the following inequality is satisfied

$$\mathbb{E}_{P}\mathbb{E}_{Q}(l_{0-1}(h(\mathbf{x}), \mathbf{y})) \le \mathbb{E}_{\hat{P}_{n}}\mathbb{E}_{Q}(l_{0-1}(h(\mathbf{x}), \mathbf{y})) + \sqrt{\frac{D(Q||P) + \log(1/\delta) + \frac{5}{2}\log n + 8}{2n - 1}}$$
(19)

A direct application of this gives us that with probability at least $1 - \omega$:

$$\mathbb{E}_{w \sim \delta} R(h, \eta, w) \leq \mathbb{E}_{w \sim \delta} \left[\frac{1}{\alpha_0} \langle l(h) - \eta, w \rangle_{\hat{P}_n} \right] + \eta + \sqrt{\frac{\mathrm{KL}(\delta \mid\mid \pi) + \log(1/\omega) + \frac{5}{2}\log n + 8}{2n - 1}}$$

E31 Let $\hat{R}_D(h, \eta, w) = \frac{1}{\alpha_0} \langle l(h) - \eta, w \rangle_{\hat{P}_n} + \eta$ Since the above inequality holds for any data dependent 632 δ :.

$$\sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} R(h,\eta,w) \le \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \left[\hat{R}_D(h,\eta,w) + \eta + \sqrt{\frac{\mathrm{KL}(\delta \mid\mid \pi) + \log(1/\omega) + \frac{5}{2}\log n + 8}{2n - 1}} \right]$$

Further, we make use of the fact $KL(\delta \parallel \pi) \leq \gamma$.

$$\leq \sup_{\delta_1 \in \Delta(\mathcal{W},\gamma)} \left[\hat{R}_D(h,\eta,w) \right] + \sup_{\delta_2 \in \Delta(\mathcal{W},\gamma)} \left[\sqrt{\frac{\operatorname{KL}(\delta_2 \mid\mid \pi) + \log(1/\omega) + \frac{5}{2}\log n + 8}{2n - 1}} \right]$$

634 Thus,

$$\sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} R(h,\eta,w) - \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D(h,\eta,w) \le \left[\sqrt{\frac{\gamma + \log(1/\delta) + \frac{5}{2}\log n + 8}{2n - 1}} \right]$$

To actually apply this uniformly over h, η , we would first need two sided concentration which we derive below as follows:

Let $a_i = \hat{R}_D(h, \eta, \delta) - R(h, \eta, \delta)$, Since $R(h, \eta, \delta) \le M/\alpha_0$, we can apply Hoeffding bound with $t = \lambda/n$ in Lemma G.1 on a_i to get:

$$\mathbb{E}_{\mathcal{D}}\exp\left(\lambda \cdot a_{i}\right) \leq \exp\frac{\lambda^{2}(M/\alpha_{0})^{2}}{8n} \mathbb{E}_{\pi}\mathbb{E}_{\mathcal{D}}\exp\left(\lambda \cdot a_{i}\right) \leq \mathbb{E}_{\pi}\exp\frac{\lambda^{2}(M/\alpha_{0})^{2}}{8n}$$

Applying Fubini's Theorem, followed by the Donsker Varadhan variational formulation we get:

$$\mathbb{E}_{\mathcal{D}}\mathbb{E}_{\pi}\left[\exp\left(\lambda\cdot a_{i}\right)\right] \leq \mathbb{E}_{\pi}\exp\frac{\lambda^{2}(M/\alpha_{0})^{2}}{8n}$$
$$= \mathbb{E}_{\mathcal{D}}\exp\sup_{\delta\in\Delta(\mathcal{W},\gamma)}\left[(\lambda\cdot a_{i}) - \mathrm{KL}(\delta\mid\mid\pi)\right] \leq \exp\frac{\lambda^{2}(M/\alpha_{0})^{2}}{8n}$$

640 The Chernoff bound finally gives us with probability $\geq 1 - \omega$:

$$\mathbb{E}_{\hat{P}_n} \mathbb{E}_Q((h(\mathbf{x}), \mathbf{y})) \lesssim \mathbb{E}_P \mathbb{E}_Q((h(\mathbf{x}), \mathbf{y})) + \frac{M}{\alpha_0} \sqrt{\frac{\mathrm{KL}(\delta \mid\mid \pi) + \log(1/\omega) + \log n}{2n - 1}}$$

Using the reverse form of the empirical PAC Bayes inequality, we can do a derivation similar to the one following the PAC-Bayes bound in Lemma G.4 to get for any fixed $\eta \in [0, M]$, $h \in \mathcal{H}$ we get:

$$\begin{aligned} \left| \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} R(h,\eta,w) - \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D(h,\eta,w) \right| &\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\mathrm{KL}(\delta \mid\mid \pi) + \log(1/\omega) + \log n}{2n - 1}} \\ &\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log(1/\omega) + \log n}{2n - 1}} \end{aligned}$$

Because we see that in the above bound the dependence on δ , is given by a log term we are essentially getting an "exponential-like" concentration. So we can think about applying uniform convergence bounds over the class $\mathcal{H} \times [0, M]$ to bounds the above with high probability $\forall (h, \eta)$ pairs.

We will now try to get uniform convergence bounds with two approaches that make different assumptions on the class of functions l(h). The first is very generic and we will show why such a generic assumption is not sufficient to get an upper bound on the generalization that is $O(1/\sqrt{n})$ in the worst case. Then, in the second approach we show how assuming a parameterization will fetch us a rate of that form if we additionally assume that the loss function is *L*-Lipschitz.

651 Approach 1:

Assume l(h) lies in a class of $(\alpha, 1)$ -Hölder continuous functions Now we shall use the following covering number bound for $(\alpha, 1)$ -Hölder continuous functions to get a uniform convergence bound over $\mathcal{H} \times [0, M]$.

Lemma G.5 (Covering number $(\alpha, 1)$ -Hölder continuous). Let \mathcal{X} be a bounded convex subset of \mathbb{R}^d with non-empty interior. Then, there exists a constant K depending only on α and d such that

$$\log \mathcal{N}(\epsilon, C_1^{\alpha}(\mathcal{X}), \|\cdot\|_{\infty}) \le K\lambda(\mathcal{X}^1) \left(\frac{1}{\epsilon}\right)^{d/\alpha}$$
(20)

for every $\epsilon > 0$, where $\lambda(\mathcal{X}^1)$ is the Lebesgue measure of the set $\{x : ||x - \mathcal{X}|| \le 1\}$. Here, $C_1^{\alpha}(\mathcal{X})$ refers to the class of $(\alpha, 1)$ -Hölder continuous functions. We assume that l(h) is $(\alpha, 1)$ -Hölder continuous. And therefore by definition, of $R(h, \eta, \cdot)$, the function is $(\alpha, 1)$ -Hölder continuous in $(l(h), \eta)$. Similat argument applies for $\sup_{\delta \in \Delta(W, \gamma)} \mathbb{E}_{w \sim \delta} R(h, \eta, w)$ since taking a pointwise supremum for a linear function over a convex set $\Delta(W, \gamma)$ would retain Hölder continuity for some value of α . Applying the above we get:

$$\log \mathcal{N}(\epsilon, \sup_{\delta \in \Delta(\mathcal{W}, \gamma)} \mathbb{E}_{w \sim \delta} R(\cdot, \cdot, w), \|\cdot\|_{\infty}) \lesssim \left(\frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log(1/\omega) + \log n}{2n - 1}}\right)^{-(d/\alpha)}$$

Now, we can show that with probability at least $1 - \delta$, $\forall h \in \mathcal{H}$ we get:

$$\left|\sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} R(h,\eta,w) - \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D(h,\eta,w)\right|$$
(21)

$$\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log(\mathcal{N}(\epsilon, R(\cdot, \cdot, w), \|\cdot\|_{\infty})/\delta) + \log n}{2n - 1}}$$
(22)

$$\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + \left(\left(\frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log(1/\delta) + \log n}{2n-1}}\right)^{-(d/\alpha)}\right) + \log(1/\delta) + \log n}{2n-1}} \qquad (23)$$

Note that in the above bound we cannot see if this upper bound shrinks as $n \to \infty$, without assuming something very strong about α . Thus, we need covering number bounds that do not grow exponentially with the input dimension. And for this we turn to parameterized classes, which is the next approach we take. It is more for the convenience of analysis that we introduce the following parameterization.

669 Approach 2:

Let $l(\cdot, \cdot)$ be a [0, M] bounded *L*-Lipschitz function in $\|\cdot\|_2$ over Θ where \mathcal{H} be parameterized by a convex subset $\Theta \subset \mathbb{R}^d$. Thus we need to get a covering of the loss function $\sup_{\delta} \mathbb{E}_{w \sim \delta} R(\theta, \eta, w)$ in $\|\cdot\|_{\infty}$ norm, for a radius ϵ . A standard practice is to bound this with a covering $\mathcal{N}(\Theta, \frac{\epsilon}{L}, \|\cdot\|_2)$, where $\|\cdot\|_2$ is Euclidean norm defined on $\Theta \subset \mathbb{R}^d$.

Lemma G.6 (Covering number for $\mathcal{N}(\Theta \times [0, M], \frac{\epsilon}{L}, \|\cdot\|_2)$ [55]). Let Θ be a bounded convex subset of \mathbb{R}^d with.

$$\mathcal{N}(\epsilon/L,\Theta,\|\cdot\|) \lesssim \left(1+\frac{L}{\epsilon}\right)^{d+1}$$
 (24)

ı.

⁶⁷⁶ We now re-iterate the steps we took previously:

T

$$\sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} R(h,\eta,w) - \sup_{\delta \in \Delta(\mathcal{W},\gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D(h,\eta,w)$$
(25)

$$\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log(\mathcal{N}(\epsilon, R(\cdot, \cdot, w), \|\cdot\|_{\infty})/\delta) + \log n}{2n - 1}}$$
(26)

$$\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + \log\left(1 + \frac{L}{\sqrt{\gamma/n}}\right)^{d+1} + \log(1/\delta) + \log n}{2n - 1}} \tag{27}$$

$$\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + (d+1)\log\left(\frac{L^2n}{\gamma}\right) + \log(1/\delta) + \log n}{2n - 1}}$$
(28)

Note that the above holds with probability at least $1 - \delta$ and for $\forall h, \eta$. Thus, we can apply it twice:

$$\begin{split} \left| \sup_{\delta \in \Delta(\mathcal{W}, \gamma)} \mathbb{E}_{w \sim \delta} R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w) - \sup_{\delta \in \Delta(\mathcal{W}, \gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w) \right| \\ &\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + (d+1)\log\left(\frac{L^2 n}{\gamma}\right) + \log(1/\delta) + \log n}{2n - 1}} \\ \left| \sup_{\delta \in \Delta(\mathcal{W}, \gamma)} \mathbb{E}_{w \sim \delta} R(h^*, \eta^*, w) - \sup_{\delta \in \Delta(\mathcal{W}, \gamma)} \mathbb{E}_{w \sim \delta} \hat{R}_D R(h^*, \eta^*, w) \right| \\ &\lesssim \frac{M}{\alpha_0} \sqrt{\frac{\gamma + (d+1)\log\left(\frac{L^2 n}{\gamma}\right) + \log(1/\delta) + \log n}{2n - 1}} \end{split}$$

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where h^* , η^* are the optimal for \mathcal{L}^*_{cvar} . Combining the two above proves the statement in Theorem B.1.

680 G.4 Proof of Theorem B.2

Setup. Let us assume there exists a prior Π such that $\mathcal{W}(\gamma)$ in Definition 2.4 is given by an RKHS induced by Mercer kernel $k : \mathcal{X} \times \mathcal{X} \mapsto \mathbb{R}$, s.t. the eigenvalues of the kernel operator decay polynomially:

$$\mu_j \lesssim j^{-2/\gamma}$$

for $(\gamma < 2)$. We solve for \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$ by doing kernel ridge regression over norm bounded $(||f||_{W(\gamma)} \le M)$ smooth functions f. Thus, $W(\gamma)$ is compact.

$$\underset{w \in \mathcal{W}(\gamma)), \|w\|_{\mathcal{W}(\gamma)} \le R}{\operatorname{arg\,max}} \quad R(h, \eta, w) = \underset{w \in \mathcal{W}(\gamma)), \|w\|_{\mathcal{W}(\gamma)} \le R}{\operatorname{arg\,max}} \quad \langle l(h) - \eta, w \rangle_P + \eta$$
(29)

$$\underset{w \in \mathcal{W}(\gamma), \|w\|_{\mathcal{W}(\gamma)} \le R}{\operatorname{arg\,max}} \quad \mathbb{E}_P \mathbb{1}((l(h) - \eta) \cdot w > 0) \tag{30}$$

We show that we can control: (i) the pessimism of the learned solution; and (ii) the generalization error (Theorem B.2). Formally, we refer to pessimism for estimates \hat{h}_D^{γ} , $\hat{\eta}_D^{\gamma}$:

excess risk or pessimism:
$$\sup_{w \in \mathcal{W}(\gamma)} |\inf_{h,\eta} R(h,\eta,w) - R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w)|$$

Theorem G.7 ((restated for convenience) bounded RKHS). For l, \mathcal{H} in Theorem B.1, and for $\mathcal{W}(\gamma)$ described above $\exists \gamma_0 \ s.t.$ for all sufficiently bitrate-constrained $\mathcal{W}(\gamma)$ i.e., $\gamma \leq \gamma_0$, w.h.p. $1 - \delta$ worst risk generalization error is $\mathcal{O}\left((1/n)\left(\log(1/\delta) + (d+1)\log(nR^{-\gamma}L^{\gamma/2})\right)\right)$ and the excess risk is $\mathcal{O}(M)$ for $\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}$ above.

692 Generalization error proof:

⁶⁹³ Note that the objective in equation 30 is a non-parametric classification problem. We can convert this

to the following non-parametric regression problem, after replacing the expectation with plug-in \hat{P}_n .

$$\inf_{w \in \mathcal{W}(\gamma)), \|w\|_{\mathcal{W}(\gamma)} \le R} \frac{1}{n} \sum_{i=1}^{n} (w(x_i, y_i) - (l(h(x_i), y_i) - \eta) + \epsilon_i)^2 + \lambda_n \|w\|_{\mathcal{W}(\gamma)}^2$$
(31)

where $\lambda_n \to 0$ as $n \to \infty$. Essentially, for non-parametric kernel ridge regression regression the regularization can be controlled to scale with the critical radius, that would give us better estimates and tighter localization bounds as we will see.

Note that in the above problem we add variable ϵ_i which represents random noise $\sim \mathcal{N}(0, \sigma_2)$. Let $\sigma_2 = 1$ for convenience. Since the noise is zero mean and random, any estimator maximizing the above objective on \hat{P}_n would be consistent with the estimator that has a noise free version. We can also thing of this as a form regularization (similar to λ), if we consider the kernel ridge regression problem as the means to obtain the Bayesian predictive posterior under a Bayesian prior that is a Gaussian Process $\mathcal{GP}(\mathbf{0}, \sigma_2 \mathbf{k}(\mathbf{x}, \mathbf{x}))$, under the same kernel as defined above.

⁷⁰⁴ First we will show estimation error bounds for the following KRR estimate:

$$\hat{w}_{D}^{\gamma} = \operatorname*{arg\,min}_{w \in \mathcal{W}(\gamma)), \|w\|_{\mathcal{W}(\gamma)} \le R} \frac{1}{n} \sum_{i=1}^{n} (w(x_{i}, y_{i}) - (l(h(x_{i}), y_{i}) - \eta) + \epsilon_{i})^{2} + \lambda_{n} \|w\|_{\mathcal{W}(\gamma)}^{2}$$
(32)

The estimation error would be measured in terms of \hat{P}_n norm *i.e.*, $\|\hat{w}_D^{\gamma} - w^*\|_{\hat{P}_n}$ where

$$w_*^{\gamma}(x,y) = \arg\min_{w \in \mathcal{W}(\gamma)), \|w\|_{\mathcal{W}(\gamma)} \le R} \quad \mathbb{E}_P \mathbb{E}_{\epsilon}((l(h(x),y) - \eta) - w(x,y) + \epsilon)^2$$

⁷⁰⁶ is the best solution to the optimization objective in population.

707 Next steps:

• First, we get the estimation error in $\|\hat{w}_D^{\gamma} - w_*^{\gamma}\|_{\hat{P}_n}$ of \hat{P}_n .

• Then using uniform laws [55] we can extend it to $L^2(P)$ norm *i.e.*, $\|\hat{w}_D^{\gamma} - w^*\|_p$.

• Then we shall prove that if we convert the \hat{w}_D^{γ} and w^* into prediction rules: $\hat{w}_D^{\gamma} \ge 0$ and w_*^{γ} , then we can get the estimation error of predictor $\hat{w}_D^{\gamma} \ge 0$ with respect to the optimal decision rule $w_*^{\gamma} \ge 0$ in class $\mathcal{W}(\gamma)$.

• The final step would give us an oracle inequality of the form in Theorem B.1.

Based on the outline above, let us start with getting $\|\hat{w}_D^{\gamma} - w^*\|_{\hat{P}_n}$. For this we shall use concentration inequalities from localization bounds (see Lemma G.8). Before we use that, we define the quantity δ_n , which is the critical radius (see Ch. 13.4 in [55]). For convenience, we also state it here. Formally, δ_n is the smallest value of δ that satisfies the following inequality (critical condition):

$$\frac{\mathcal{R}_n(\delta)}{\delta} \le \frac{R}{2} \cdot \delta \tag{33}$$

718 where,

$$\mathcal{R}_n(\delta) \coloneqq \mathbb{E}_{\epsilon} \left[\sup_{g \in (\mathcal{F} - f^*), \|g\|_{\mathcal{F}} \leq R, \|g\|_{\hat{P}_n} \leq \delta} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i g(x_i, y_i) \cdot l(h(x_i) - y_i) \right| \right]$$

and ϵ is some sub-Gaussian zero mean random variable.

Lemma G.8 ([55]). For some convex RKHS class \mathcal{F} Let \hat{f} be defined as:

$$\hat{f} \in \operatorname*{arg\,min}_{f \in \mathcal{F}, \|f\|_{\mathcal{F}} \le R} \left\{ \frac{1}{n} \sum_{i=1}^{n} (y_i - f(x_i))^2 + \lambda_n \|f\|_{\mathcal{F}}^2 \right\}$$

then, with probability $\geq 1 - c_2 \exp\left(-c_3 \frac{nR^2 \delta_n^2}{\sigma^2}\right)$ and when $\lambda_n \geq \delta_n^2$ we get:

$$\|\hat{f} - f^*\|_2^2 \le c_0 \inf_{\|\mathcal{F}\| \le R} \|f - f^*\|_n^2 + c_1 R^2 (\delta_n^2 + \lambda_n).$$

722 Note that it is standard exercise in statistics to derive the following closed form for the problem in 723 equation 32:

$$\hat{w}_D^{\gamma}(\cdot) = \hat{K}_n(\cdot, Z)(\hat{K}_n^T \hat{K}_n + \lambda_n I)^{-1} \left(l(h)_D - \epsilon_D \right)$$

where $l(h)_D$ is the loss vector and ϵ_D is the noise vector for dataset \mathcal{D} and \hat{K}_n is the empirical kernel matrix given by $\hat{K}_n = \frac{1}{2} h((x, y_n))$ and Z is a matrix of (x, y) pairs in dataset

matrix given by $\hat{K}_{i,j} = \frac{1}{n}k((x_i, y_i), (x_j, y_j))$, and Z is a matrix of (x, y) pairs in dataset.

Corollary G.9. [55] Let $\hat{\mu}_j$ be the eigen values $\hat{\mu}_1 \geq \hat{\mu}_2 \dots \geq \hat{\mu}_n$ for the empirical Kernel matrix

727 \hat{K} , then we have for any δ satisfying

$$\left| \frac{2}{n} \left(\sum_{i=1}^{n} \min(\delta^2, \hat{u}_j) \right) \right| \leq \frac{R}{4} \delta^2$$

, it is necessary that δ satisfies the critical condition in equation 33.

1

To show the above critical condition we shall now use the polynomial decaying property that for the specific kernel induced by $W(\gamma)$, as stated in our assumption in the beginning of this section. For this we take standard approach taken for polynomial decay kernels [61]. Let $\exists C$ for some large C > 0

rse such that $\hat{\mu}_j \leq C j^{-2/\gamma}$. Then for some k, such that $\delta^2 \geq c k^{-2/\gamma}$

$$\sqrt{\frac{1}{n} \left(\sum_{j=1}^{n} \min(\delta^{2}, \hat{\mu}_{j})\right)} \lesssim \sqrt{\frac{2}{n} \left(\sum_{i=1}^{n} \min(\delta^{2}, Cj^{-2/\gamma})\right)} \\
\lesssim \sqrt{\frac{2}{n} \left(k\delta^{2} + C\sum_{j=k+1}^{n} j^{-2/\gamma}\right)} \lesssim \sqrt{\frac{2}{n} \left(k\delta^{2} + C\sum_{j=k+1}^{\infty} j^{-2/\gamma}\right)} \\
\lesssim \sqrt{\frac{2}{n} \left(k\delta^{2} + C\int_{j=k+1}^{\infty} z^{-2/\gamma} dz\right)} \lesssim \sqrt{\frac{2}{n} \left(k\delta^{2} + Ck^{-2/\gamma+1} dz\right)} \\
\le \sqrt{2/n} \left(\sqrt{k} \cdot \delta\right) \le \frac{1}{\sqrt{n}} \cdot \delta^{1-\gamma/2}$$

⁷³³ Now, setting the above into the critical condition equation from Corollary above:

$$\frac{1}{\sqrt{n}} \cdot \delta^{1-\gamma/2} \le \frac{R}{4} \delta^2$$

$$\implies \delta^{1+\gamma/2} \ge \frac{1}{\sqrt{nR}}$$

734 This tells us that:

$$\delta_n^2 \gtrsim \left(\frac{1}{nR^2}^{\frac{2}{\gamma+2}}\right) \tag{34}$$

735 is the critical radius.

We shall later plug this into the bound we have into a uniform bound over the concentration inequality in Lemma G.8. The reason we need a uniform bound over Lemma G.8 is that in its current form, it only bounds $\|\hat{w}_D^{\gamma} - w_*^{\gamma}\|_{\hat{P}_n}^2$ for a specific choice of η, h . In order to arrive at the worst risk generalization error of the form we have in Theorem B.1 we need to satisfy that with high probability $1 - \delta \forall \eta, h$, a critical concentration bound of the form in Lemma G.8 but over $\sup_{\eta,h} \|\hat{w}_D^{\gamma} - w_*^{\gamma}\|_{\hat{P}_n}^2$.

Let $\epsilon = c_2 \exp\left(c_3 n R^2 \frac{\delta_n^2}{\sigma^2}\right)$. Since δ_n^2 needs to be large enough (see condition in equation 34), we use Lemma G.8 in the following bound, incorporating δ_n condition we derived. 743 With high probability $1 - \epsilon$:

$$\|\hat{w}_{D}^{\gamma} - w_{*}^{\gamma}\|_{\hat{P}_{n}}^{2} \lesssim \inf_{w \in \mathcal{W}(\gamma), \|w\| \leq R} \|w - w_{*}^{\gamma}\|_{\hat{P}_{n}}^{2} + R^{2} \max\left(\left(\frac{1}{nR^{2}}\right)^{\frac{\gamma+2}{\gamma}}, \left(\log(1/\epsilon)\frac{1}{nR^{2}}\right)\right)$$
(35)

To apply uniform convergence argument on the above we would need to apply a union bound on a covering of $\Theta \times [0, M]$, so that we get the probability bound to hold for all η, h .

For this we use the same technique as in the proof of Theorem B.1. First, we shall use Lemma G.6 to get a covering number bound for bounded convex subset Θ of \mathbb{R}^d that parameterizes the learner (Theorem B.2).

$$\mathcal{N}(\beta/L, \Theta \times [0, M], \|\cdot\|) \lesssim \left(1 + \frac{L}{\beta}\right)^{d+1}$$
(36)

And we know that a covering of $\Theta \times [0, M]$ in radius β/L , will fetch a covering for $l(h) - \eta$ in β , since we assume $l(\cdot)$ to be Lipschitz in θ . Thus, all we need to prove bound equation 35 holds uniformly is to get a covering in radius $R^2 \max\left(\left(\frac{1}{nR^2}\right)^{\frac{2}{\gamma+2}}, \left(\log(1/\epsilon)\frac{1}{nR^2}\right)\right)$. Thus, a acovering in $R^2 = R^2\left(\left(\frac{1}{nR^2}\right)^{\frac{2}{\gamma+2}}\right)$. Thus, the number of elements in cover are:

$$J = \left(1 + \frac{L}{\left(R^2\left(\frac{1}{nR^2}^{\frac{2}{\gamma+2}}\right)\right)}\right)^{d+1}$$

753 For union bound we need:

$$\begin{aligned} J\epsilon/c_2 &= \exp\left(-c_3 n R^2 \delta_n^2\right) \\ \implies \log\left(\frac{1}{\epsilon}\right) + \log J \gtrsim c_3 n R^2 \delta_n^2 \\ \implies \log\left(\frac{1}{\epsilon}\right) + (d+1) \log\left(\frac{L}{\left(R^2\left(\left(\frac{1}{nR^2}\right)^{\frac{2}{2+\gamma}}\right)\right)}\right) \gtrsim c_3 n R^2 \delta_n^2 \\ \implies \log\left(\frac{1}{\epsilon}\right) + (d+1) \log\left(\left(LR^{-2}\right)^{\frac{\gamma+2}{2}} n R^2\right) \gtrsim c_3 n R^2 \delta_n^2 \end{aligned}$$

The uniform convergence bound that we get is $R^2 \max\left(\left(\frac{1}{nR^2}\right)^{\frac{\gamma+2}{\gamma}}, \left(\log(J/\epsilon)\frac{1}{nR^2}\right)\right)$. In the above sequence of steps we have shown that, due to the size of J, the second term would be maximum, or at least there exists a γ_0 , such that the second term would be higher for all $\gamma \ge \gamma_0$, for any sample size.

Thus, we get the following probabilistic uniform convergence. With probability $\geq 1 - \epsilon, \forall \eta, h:$

$$\|\hat{w}_{D}^{\gamma} - w_{*}^{\gamma}\|_{\hat{P}_{n}}^{2} \lesssim \inf_{w \in \mathcal{W}(\gamma), \|w\| \le R} \|w - w_{*}^{\gamma}\|_{\hat{P}_{n}}^{2}$$
(37)

$$\lesssim \frac{1}{n} \log(\frac{1}{\epsilon}) + (d+1) \log\left(\left(LR^{-2}\right)^{\frac{\gamma+2}{2}} nR^2\right)$$
(38)

$$\lesssim \frac{1}{n} \log(\frac{1}{\epsilon}) + (d+1) \log\left(\left(L^{\gamma/2} R^{-\gamma}\right) n\right)$$
(39)

(40)

- Applying the above twice, once one \hat{w}_D^{γ} and another on w_*^{γ} we prove the generalization bound in 758
- Theorem B.2. 759
- Excess risk bound: 760
- In the same setting we shall now prove the excess risk bound. Recall the definition of excess risk: 761

excess risk :=
$$\sup_{w \in \mathcal{W}(\gamma)} |\inf_{h,\eta} R(h,\eta,w) - R(\hat{h}_D^{\gamma},\hat{\eta}_D^{\gamma},w)|.$$

762 Let $h^*(w), \eta^*(w) = \inf_{h,\eta} R(h, \eta, w)$, then:

excess risk =
$$\sup_{w \in \mathcal{W}(\gamma)} |\inf_{h,\eta} R(h,\eta,w) - R(\hat{h}_D^{\gamma},\hat{\eta}_D^{\gamma},w)|$$
(41)

$$\leq \sup_{w \in \mathcal{W}(\gamma)} \left(R(h^*(w), \eta^*(w), w) - R(\hat{h}_D^{\gamma}, \hat{\eta}_D^{\gamma}, w) \right)$$
(42)

$$\leq \sup_{w \in \mathcal{W}(\gamma)} \left(\frac{1}{\alpha_0} \langle l(h^*) - l(\hat{h}_D^{\gamma}) - (\eta^*(w) - \hat{\eta}_D^{\gamma}), w \rangle_P \right)$$
(43)

$$\leq \frac{M}{\alpha_0} \sup_{w \in \mathcal{W}(\gamma)} \left((\|w\|_{L_2(P)}) \right) \tag{44}$$

- Note, that according to our assumption $||w||_{\mathcal{W}(\gamma)} \leq B$ *i.e.*, the smooth functions are bounded in 763
- RKHS norm. The following lemma relates bounds in RKHS norm to bound in $L_2(P)$ bound for 764 kernels with bounded operator norms: 765
- **Lemma G.10.** For an RKHS \mathcal{H}_k with norm $\|\cdot\|_{\mathcal{H}_k}$: 766

$$\|f\|_{L^{2}(P)} = \|T_{K}^{1/2}f\|_{\mathcal{H}_{k}} \leq \sqrt{\|T_{K}^{1/2}\|_{op}} \cdot \|f\|_{\mathcal{H}_{k}}$$

Proof: 767

$$\begin{aligned} \|T_K^{1/2}f\|_{\mathcal{H}_k}^2 &= \langle T_K^{1/2}f, T_K^{1/2}f \rangle_{\mathcal{H}_k} = \langle f, T_Kf \rangle_{\mathcal{H}_k} \\ &= \sum_{j=1}^{\infty} \frac{\langle \phi_j, f \rangle_{L^2(P)}, \langle \phi_j, T_Kf \rangle_{L^2(P)}}{\lambda_j} \\ &= \|f\|_{L^2(P)}^2 \end{aligned}$$

- 768
- In the above λ_j are the Eigen values of the kernel and the Eigen functions ϕ_j are orthonormal and span $L^2(P)$ / Thus, $\|f\|_{L^2(P)} \leq \|T_K^{1/2}\|_{op} \cdot \|f\|_{\mathcal{H}_k}$. Since we assume polynomially decaying Eigen values for our kernel, it is easy to see that $\|T_K^{1/2}\|_{op} = \mathcal{O}(1)$. 769 770
- Applying Lemma G.10 to equation 44, directly gives us the excess risk bound and completes the 771 proof. 772

excess risk
$$\lesssim \|T_K^{1/2}\|_{op} \cdot B = \mathcal{O}(B)$$

G.5 Proof of Theorem B.4 773

Setup. The algorithm is as follows: Consider a two-player zero-sum game where the learner uses a 774 no-regret strategy to first play $h \in \mathcal{H}, \eta \in \mathbb{R}$ to minimize $\mathbb{E}_{w \sim \delta} R(h, \eta, w)$. Then, the adversary plays 775

- follow the regularized leader (FTRL) strategy to pick distribution $\delta \in \Delta(W, \gamma)$ to maximize the same.
- The regularizer used is a negative entropy regularizer. Our goal is to analyze the bitrate-constraint γ 's effect on the above algorithm's convergence rate and the pessimistic nature of the solution found. For
- effect on the above algorithm's convergence rate and the pessimistic nature of the solution found. For this, we need to first characterize the bitrate-constraint class $W(\gamma)$. So we assume there exists a prior
- ⁷⁸⁰ If such that $\mathcal{W}(\gamma)$ is Vapnik-Chervenokis (VC) class of dimension $O(\gamma)$.
- Note that $R(h, \eta, w)$ is convex in h and linear in η, l . Thus, as we discuss in the derivation for equation 6 this objective optimized over convex sets has a unique saddle point (Nash equilibrium) by Weierstrass's theorem. Thus, to avoid repetition we only discuss the proofs for the other two claims
- 784 on convergence and excess risk.
- 785 Convergence:

Given that $W(\gamma)$ is a VC class of dimension $C\gamma$ for some large C, we can use Sauer-Shelah [6] Lemma (stated) below to bound the total number of groups that can be identified by $W(\gamma)$ in npoints.

Lemma G.11 (Sauer's Lemma). *The Vapnik-Chervonenkis dimension of a class* \mathcal{F} *, denoted as VC-dim*(\mathcal{F})*, and it is the cardinality of the largest set S shattered by* \mathcal{F} *. Let* $d = VC - dim(\mathcal{F})$ *, then*

for all $m, C[m] = \mathcal{O}(m^d)$

Thus, the total number of groups that can be proposed on n points by $W(\gamma)$ is $\mathcal{O}(n^{\gamma})$. A similar observation was made in Kearns et al. [27]. Different from them, our goal is to analyze the algorithm iterates for our solver described above and bound its pessimism.

First, for convergence rate we show that the above algorithm has a low regret—a standard exercise in online convex optimization. Note that any distribution picked by the adversary can be seen as multinomial over a finite set of possible groups that is let's say K, and from discussion above we

know that $K = O(n^{\gamma})$. Further, the negative entropy regularizer is given as:

$$B(\delta) \coloneqq c \cdot \sum_{i=1}^{K} \delta_i \log \delta_i \tag{45}$$

where the sum is over total possible groups identified by $W(\gamma)$. Let the probability assigned to group *i* be denoted as δ_i . The FTRL strategy for adversary is given as:

$$\delta_T = \operatorname*{arg\,min}_{\delta \in \Delta(\mathcal{W}(\gamma))} \sum_{t=1}^{T-1} \frac{1}{\alpha_0} \langle l(h_t) - \eta_t, \delta_t \rangle_{\hat{P}_n} + \eta + c \cdot \sum_{i=1}^K \delta_i \log \delta_i$$
(46)

Then the regret for not having picked a single action δ is given as:

$$\operatorname{REGRET}_{T}(\delta) \coloneqq \sum_{t=1}^{T} \frac{1}{\alpha_{0}} \langle l(h_{t}) - \eta_{t}, \delta_{t} - \delta_{t+1} \rangle_{\hat{P}_{n}} + B(\delta) - B(\delta_{1})$$
(47)

We bound the two terms in the above bound separately. With $\sum_{k=1}^{k} \delta_k = 1$, we get the strong dual for the FTRL update above as:

$$\sum_{t=1}^{T-1} \frac{1}{\alpha_0} \langle l(h_t) - \eta_t, \delta_t \rangle_{\hat{P}_n} + \eta + c \cdot \sum_{i=1}^K \delta_i \log \delta_i + \lambda \cdot (\sum_{i=1}^K \delta_i - 1)$$
(48)

804 Solving we get:

$$\delta_t(k) = \frac{\exp\left(\frac{-1}{c}\right) \sum_{t=1}^{t-1} \mathbb{E}_{\hat{P}_n} \frac{1}{\alpha_0} (l(h_t) - \eta_t | G_k) + \eta/K}{\sum_{k=1}^K \exp\left(\frac{1}{\alpha_0} \frac{-1}{c}\right) \sum_{k=1}^{t-1} (\mathbb{E}_{\hat{P}_n} \frac{1}{\alpha_0} (l(h_t) - \eta_t | G_k) + \eta/K)}$$
(49)

where $\mathbb{E}_{\hat{P}_n}(l(h_t) - \eta_t | G_k)$ is the expected empirical loss in group G_k and $\delta_t(k)$ is the adversary's distribution at time step t for the k^{th} group.

807 Claim on stability:

$$\frac{1}{\alpha_0} \langle l(h_t) - \eta_t, \delta_t - \delta_{t+1} \rangle_{\hat{P}_n} \le 1/c$$
(50)

808 The above statement is true because,

$$\delta_{t+1}(i) = \delta_t(i) \cdot \exp\left(\frac{1}{\alpha_0 c} \mathbb{E}[l(h_t) - \eta_t | G_i] + \eta_t / K\right)$$
(51)

Thus, if $l(h_t) \in [0, M/\alpha_0]$, *i.e.*, losses are bounded then:

$$\delta_{t+1}(i) \ge \delta_t(i) \cdot e^{-1/c} \ge \delta_t(i) \cdot (1 - 1/c).$$
(52)

and our stability claim is easy to see. Thus, we have bounded the first term in our regret bound above. Further, we can to see that $B(x) - B(x_1) \le c \log K$. Thus, we have bounded both terms in the regret

⁸¹² bound above in terms of
$$c$$

$$\operatorname{REGRET}_T \le (T/c) + (c \log K) \tag{53}$$

813 Setting $c = \sqrt{\frac{T}{\log K}}$, we get:

$$\frac{\operatorname{REGRET}_T}{T} \le \sqrt{\frac{\log K}{T}}$$
(54)

Now, our VC claim gave $K = \mathcal{O}(n^{\gamma})$. Hence,

$$\frac{\text{REGRET}_T}{T} = \mathcal{O}\sqrt{\frac{\gamma \log n}{T}}$$
(55)

- Next, we use Theorem 9 from Abernethy et al. [1] that maps low regret $O(\epsilon)$ algorithms in zero-sum
- s16 convex-concave games to ϵ -optimal equilibriums.
- ⁸¹⁷ Let regret be ϵ , then applying their theorem gives us:

$$V^* - \epsilon \le \inf_{h \in \mathcal{H}, \eta \in \mathbb{R}} R_D(h, \eta, \bar{\delta}_T) \le V^* \le \sup_{\delta \in \Delta(\mathcal{W}(\gamma))} R_D(\bar{h}_T, \bar{\eta}_T, \delta) \le V^* + \epsilon$$
(56)

818 where

$$V^* = R_D(h_D^*(\gamma), \eta_D^*(\gamma), \delta_D^*(\gamma)) = \inf_{h \in \mathcal{H}, \eta \in \mathbb{R}} \sup_{\delta \in \Delta(\mathcal{W}(\gamma))} \frac{1}{\alpha_0} \langle l(h) - \eta, \delta \rangle + \eta$$
(57)

819 Excess risk:

820 For excess risk we need to bound:

$$\frac{1}{\alpha_0} \sup_{h \in \mathcal{H}, \eta \in \mathbb{R}} \left| \sup_{\delta \in \Delta(\mathcal{W}(\gamma))} \langle l(h) - \eta, \delta - \delta^*(\gamma) \rangle \right|$$
(58)

$$\leq \frac{M}{\alpha_0} \frac{1}{2} \operatorname{TV}(\delta - \delta^*(\gamma)) \leq \frac{M}{2\alpha_0} (1 - 1/K) = \frac{M}{\alpha_0} \mathcal{O}(1 - 1/n^{\gamma})$$
(59)

In the above argument we used the fact that at equilibrium, $\delta^*(\gamma)$ would be uniform over all possible distinct group assignments. This completes our proof of Theorem B.4.