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ABSTRACT

This paper studies forecasting of the future distribution of events in human action sequences, a task essential in domains like retail, finance, healthcare, and recommendation systems where the precise temporal order is often less critical than the set of outcomes. We challenge the dominant autoregressive paradigm and investigate whether explicitly modeling the future distribution or order-invariant multi-token approaches outperform order-preserving methods. We analyze local order invariance and introduce a distribution-based metric to quantify temporal drift. We find that a simple explicit distribution forecasting objective consistently surpasses complex implicit baselines. We further analyze the emergence of *mode collapse* in predicted categories, identifying and evaluating key contributing mechanisms. This work provides a principled framework for selecting modeling strategies and offers practical guidance for building more accurate and robust forecasting systems. **The code will be released upon publication.**

1 INTRODUCTION

In many real-world prediction tasks, the precise temporal ordering of events is irrelevant. Instead, predicting the distribution of outcomes, where only the presence or absence of specific elements matters, is sufficient and often more practical.

For instance, in retail operations, probabilistic demand forecasting enables optimal inventory management and supply chain planning by modeling the full range of possible product demands without requiring sequence order (Nassibi et al., 2023; Larson, 2001). Similarly, in healthcare, clinical diagnosis systems treat disease categories as unordered sets within a single hospital admission. The presence of certain conditions is clinically more significant than the exact order in which they were diagnosed (Johnson et al., 2016; Mullenbach et al., 2018). Recommendation systems further exemplify this principle known as *basket prediction* (Rendle, 2020). Finally, many multi-label problems can naturally be framed as distribution forecasting tasks.

The central focus of this paper is to model the future distribution of human actions over a fixed future horizon. In this work we consider *Event Sequences* (EvS) (Osin et al., 2025; Udovichenko et al., 2024) - temporal records of human actions which underpin a wide range of decision-making systems across domains including healthcare (Johnson et al., 2016), financial transactions (Udovichenko et al., 2024; Mollaev et al., 2024; Yang & Xu, 2019), e-commerce (Li et al., 2021), recommender systems (Shevchenko et al., 2024; Klenitskiy et al., 2024; Zhelnin et al., 2025), and human action recognition (Surkov et al., 2024). Despite its practical importance and deceptively simple formulation, distribution forecasting for EvS remains significantly understudied.

Inspired by advances in Natural Language Processing (NLP), contemporary approaches to modeling sequential behavior often default to autoregressive generation predicting the next token conditioned on an exact prefix ordering (Karpukhin et al., 2024; Klenitskiy et al., 2024). While Next Token Prediction (NTP) has long dominated sequential modeling, Multi-Token Prediction (MTP) has recently gained traction due to its demonstrated improvements in model quality and generalization, particularly in tasks such as planning, code generation and EvS forecasting (Nagarajan et al., 2025; Bachmann & Nagarajan, 2024; Yu et al., 2025; Karpukhin & Savchenko, 2024).

054
 055 **This raises a practical question:** When should we model future event distributions **explicitly**, and
 056 when is it worth preserving temporal structure through **implicit** order-preserving objectives methods
 057 like NTP or Multi-Token Prediction (MTP)? To answer this, we make the following contributions:
 058

059 **(1) We systematically study the task of forecasting the distribution of future events** over a fixed
 060 horizon and demonstrate its importance as a viable alternative to autoregressive modeling in do-
 061 mains where the order of events is weakly informative or irrelevant. Our results show that this task
 062 is not only meaningful for practical applications but also enables simpler, more robust models that
 063 avoid pitfalls such as *mode collapse* (see Sec. 4.1).
 064

065 **(2) Explicit vs. Implicit Objective Evaluation:** We conduct a rigorous empirical comparison of
 066 four training paradigms on seven public datasets: (1) Next Token Prediction , (2) Multi-Token Pre-
 067 diction with ordered output, (3) an order-invariant set prediction approach with post-hoc alignment,
 068 and (4) **GRU-Dist** - our proposed, explicit distribution forecasting objective. Our results demon-
 069 strate that explicitly modeling the future event distribution (GRU-Dist) consistently outperforms all
 070 order-preserving baselines across most domains. Surprisingly, when evaluated with order-invariant
 071 metrics, this superiority holds even on textual data, where sequential structure is traditionally as-
 072 sumed critical.
 073

074 **(3) Connecting Dataset Structure to Model Performance:** We believe that the efficacy of se-
 075 quential modeling is fundamentally governed by intrinsic dataset properties, since autoregressive
 076 paradigms developed for text we attempt to re-evaluate them accounting for dataset properties.
 077 We propose a following set of dataset characteristics and evaluations: the *Staticity Index (S)*, a
 078 distribution-based metric quantifying temporal drift across sequences; *Local Permutation Analysis*,
 079 which measures sensitivity to event shuffling within sliding windows; *Exponential Decay Parameter*
 080 λ , capturing category imbalance and *Consecutive Repeat Rate (CRR)*, a measure to analyze ration
 081 of consecutively repetitive tokens, which are present in some real world e-commerce datasets as
 082 repetitive item clicks.
 083

084 Our findings provide actionable guidance for informed model selection with respect to dataset prop-
 085 erties, and demonstrate that next-token prediction is not universally optimal, even for large models
 086 across domains.
 087

088 2 RELATED WORK

089 **Architectures for Event Sequences.** Modeling user actions sequentially by conditioning on past
 090 behavior has become an essential component of modern recommendation pipelines. These ap-
 091 proaches effectively adapt ideas from natural language processing (NLP), particularly attention-
 092 based architectures (Kang & McAuley, 2018; Sun et al., 2019; Klenitskiy et al., 2024; Mezent-
 093 sev et al., 2024). However, it remains unclear whether transformer-based architectures are in-
 094 deed the most suitable for predicting future user actions. In *EBES* Osin et al. (2025) and in *Seq-*
 095 *NAS* Udovichenko et al. (2024), the authors demonstrate that RNN-based architectures outperform
 096 transformer-based models on **EvS** classification tasks. Delving deeper into this issue, Karpukhin
 097 & Savchenko (2025) investigate the limitations of transformers and proposes several modifications
 098 that enable them to surpass RNNs in classification performance. However, as the same work further
 099 reveals, these enhancements do not translate to improved performance in forecasting future tokens.
 100 In this work, we focus on RNN- and GPT-based architectures, as they remain the most applicable in
 101 this domain.
 102

103 **Multi-Token vs. Single-Token Prediction.** Multi-Token Prediction (MTP) has recently gained
 104 traction due to its demonstrated improvements in model quality and generalization particularly in
 105 tasks such as planning, code generation (Nagarajan et al., 2025; Bachmann & Nagarajan, 2024; Yu
 106 et al., 2025). However, a key challenge lies in the common assumption that predicted tokens are
 107 conditionally independent Gloeckle et al. (2024).
 108

109 **Teacherless Learning** Bachmann & Nagarajan (2024) offers an intermediate approach between
 110 Next-Token Prediction (NTP) and MTP, conceptually opposing teacher forcing. Unlike MTP,
 111 Teacherless Learning is grounded in a rigorous mathematical framework. While it does not accel-
 112 erate inference, it addresses fundamental limitations of traditional NTP. As Nagarajan et al. (2025)
 113

108 note: “Teacherless training and diffusion models comparatively excel in producing diverse and orig-
 109 inal output.”

110 Although earlier work focused primarily on text generation, Karpukhin & Savchenko (2024) ex-
 111 tended these ideas to **EvS**, demonstrating that multi-token generation and diffusion-based ap-
 112 proaches indeed outperform the single-token paradigm. In this work, we investigate NTP, a multi-
 113 token strategy similar to that proposed in Karpukhin & Savchenko (2024) and propose a new explicit
 114 approach for distribution forecasting.

115
 116 **Order Importance in EvS.** It has been established that permuting sequences in **EvS** datasets does
 117 not degrade performance on classification tasks (Osin et al., 2025; Moskvoretskii et al., 2024), an ob-
 118 servation which significantly challenges the assumed sequential nature of this data type. Klenitskiy
 119 et al. (2024) investigates whether datasets from the domain of sequential recommender systems gen-
 120 uinely exhibit sequential structure. Specifically, the authors evaluate whether permuting sequences
 121 leads to performance degradation in next-token prediction tasks, and find that the extent of degra-
 122 dation varies by dataset, some datasets are more “sequential” than others. In this work, we extend this
 123 investigation beyond recommender systems and analyze local permutation invariance.

124 3 DATASETS

125 To evaluate the proposed methods and hypotheses, we conduct experiments on a diverse collection
 126 of real-world sequential datasets spanning multiple domains—including financial transactions, e-
 127 commerce, retail, music streaming, and literary text. A summary of key statistics is provided in
 128 Table 1; full descriptions, including preprocessing steps are available in Appendix A.3.

129
 130 Table 1: Dataset statistics and characteristics.

131 Dataset	132 ID	133 Domain	134 Sequences	135 Mean len	136 Target Field	137 Classes
Multimodal Banking Dataset 2024	MBD	Transactions	1.5M	313	Event type	55
AgeGroup Transactions	AGE	Transactions	30K	888	Small group	203
X5 RetailHero	Retail	Retail	40K	112	Level 2	43
Alphabattle-2.0	AB	Transactions	1M	213	MCC category	28
Complete Works of Shakespeare	ShS	Text	5K	106	Character	65
Megamarket (2024)	MM	E-commerce	2.73M	653	Category ID	9.8K
Zvuk (2024)	Zvuk	Music Streaming	380K	1020	Artist ID	210K
Taobao User Behavior	Taobao	E-commerce	10K	535	Item category	8K

144 4 DATASET DIAGNOSTIC

145 4.1 TEMPORAL ORDER AND MODE COLLAPSE IN EVENT SEQUENCE MODELING

146 In time series and natural language modeling, precise temporal ordering is crucial. However, in
 147 domains like system logs or bank transactions, the *exact micro-temporal order* of events within
 148 short windows may be ambiguous or irrelevant—e.g., two unrelated log entries milliseconds apart
 149 could plausibly appear in either order without changing system semantics. We illustrate this effect
 150 in Appendix 4. This motivates a formal distinction between two types of temporal structure:
 151

- 152 • **Local invariance:** Within a narrow window $W_t = (y_t, \dots, y_{t+H})$, event order is semantically
 153 irrelevant—permutations of the same multiset are equally plausible.
- 154 • **Global structure:** Across broader time intervals, dependencies between consecutive windows re-
 155 main meaningful; e.g., $p(W_2 | W_1)$ for $W_1 = (y_0, \dots, y_{t-1})$ and $W_2 = (y_t, \dots, y_{t+H})$ captures
 156 genuine temporal progression.

157 Conventional autoregressive (AR) models are trained to predict the next token y_t given its full his-
 158 tory (y_0, \dots, y_{t-1}) . To accommodate local invariance, one might relax this strict left-to-right depen-
 159 dency by defining a *prediction horizon* $\{y_t, \dots, y_{t+H}\}$ and training the model to predict *any* event

162 within this window. Under the assumption of uniform uncertainty over the horizon, the training
 163 objective becomes:
 164

$$166 \mathbb{E}_{k \sim \text{Uniform}[0, H]} [\log p(x = y_{t+k} \mid y_0, \dots, y_{t-1})] = \frac{1}{H+1} \sum_{m=0}^H \log p(x = y_{t+m} \mid y_0, \dots, y_{t-1}). \quad (1)$$

170 Critically, standard AR architectures use a *single output distribution* $q_t(\cdot)$ at time t to score all
 171 tokens in the horizon. Under local permutation invariance, the optimal q_t that maximizes the above
 172 objective is the empirical distribution over the multiset $\{y_t, \dots, y_{t+H}\}$. Consequently, the model
 173 learns a *static predictive distribution* over the entire window: $q_t \approx q_{t+1} \approx \dots \approx q_{t+H}$. This static
 174 distribution becomes problematic at inference time. When generating sequences using deterministic
 175 decoding (e.g., argmax or low-temperature sampling), the model outputs:

$$176 \hat{y}_{t+k} = \arg \max_x q_{t+k}(x) \approx \arg \max_x q_t(x), \quad \forall k \in [0, H].$$

178 Since q_t is dominated by the most frequent event in the window, the model repeatedly predicts the
 179 *empirical mode* of \mathcal{W}_t , suppressing rarer—but valid—events. We term this phenomenon *temporal*
 180 *mode collapse*.

181 We propose that explicitly modeling the distribution of events across entire windows, rather than
 182 enforcing pointwise predictions, offers a principled resolution. This allows models to better capture
 183 the stochastic nature of real-world event sequences while avoiding degenerate solutions.
 184

185 4.2 STATICITY INDEX

187 Before fitting neural models, we quantify how the event distribution of each sequence changes over
 188 time. Previous studies have shown that drift of the temporal distribution can strongly influence fore-
 189 cast performance, including context-driven shift (Chen et al., 2024), seasonality-induced shifts (Liao
 190 et al., 2025), and temporal dataset shift benchmarks (Yao et al., 2022). Motivated by these findings,
 191 we measure the stability or dynamic of the empirical event distribution within each dataset. To this
 192 end, we plot the *Shape* score drift for every dataset to reveal whether their event distributions remain
 193 nearly static or show meaningful temporal variation.

194 Several datasets contain sequences with nearly static behavior; to verify this, we plot the *Shape* score
 195 drift for each dataset.
 196

197 4.2.1 PER-FEATURE DISSIMILARITY SCORE

199 **Procedure.** For each sequence, we fix a window length W and stride s , then slide the window across
 200 the timeline. At every position i , we extract the feature distribution P_i within the current window
 201 and compare it with the baseline distribution P_0 computed from the first window. To compare them,
 202 we suggest to leverage the following score:

203 Let P_0 and P_i denote the empirical distributions in the reference window and the i -th window,
 204 respectively.

205 **Discrete features.** For categorical attributes defined on \mathcal{A} we employ the *total variation (TV) dis-
 206 tance*, $\text{TV}(P_0, P_i) = \frac{1}{2} \sum_{a \in \mathcal{A}} |P_0(a) - P_i(a)|$. Because lower TV indicates higher similarity, we
 207 report its complement $(1 - \text{TV})$, so that higher values consistently reflect better alignment.
 208

209 **Continuous features.** For numerical attributes we use the *Kolmogorov–Smirnov* statistic. Let F_0
 210 and F_i be the empirical CDFs corresponding to P_0 and P_i . The KS divergence is $\text{KS}(P_0, P_i) =$
 211 $\sup_{x \in \mathbb{R}} |F_0(x) - F_i(x)|$. Analogously, we report the similarity score $1 - \text{KS}$.

212 **Shape score.** For window i we propose to compute each feature’s distance using the appropriate
 213 formula above and then average across all features: $\text{Shape}(P_0, P_i) = \frac{1}{M} \sum_{j=1}^M d_j(P_0^{(j)}, P_i^{(j)})$,
 214 where M is the number of features and d_j is TV when the j th feature is categorical, and KS other-
 215 wise. Plotting $i \mapsto \text{shape}(P_0, P_i)$ yields the drift curves used throughout this paper.

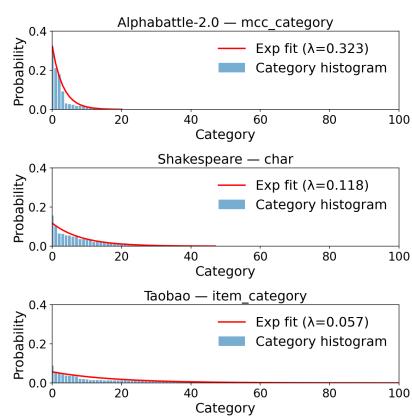


Figure 1: Distribution of categories in datasets. We present normalized number of categories.

Table 2: Dataset statistics: exponential decay parameter (λ), Consecutive Repeat Rate (CRR), total number of distinct categories (TCD), Staticity index (S ; **average distributional similarity over time, 1 = fully stationary**), and perplexity (PPL) increase after full shuffle.

Dataset	λ	CRR	TCD	S	PPL
<i>Banking domain</i>					
MBD	0.415	1.718	55	0.842	1.02 \times
AB	0.305	1.518	28	0.725	1.08 \times
Age	0.245	1.148	203	0.772	1.24 \times
Retail	0.185	1.372	43	0.782	1.27 \times
<i>Text</i>					
ShS	0.118	1.019	64	0.803	5.09 \times
<i>Recommender Systems</i>					
Taobao	0.016	4.492	1.9K	0.650	13.00 \times
MM	0.005	3.502	9.8K	0.406	10.87 \times
Zvuk	0.003	1.239	210K	0.363	5.05 \times

With these definitions we shift the window across the entire sequence and plot trajectory $i \mapsto \text{shape}(P_0, P_i)$, obtaining time-resolved drift curve that summarises how the distribution evolves over the time.

4.2.2 STATICITY IN DATASETS

Across banking datasets (MBD, Retail, Age, AlphaBattle) the majority of user sequences form static clusters with negligible temporal drift (Figure 5, Appendix A.4). In contrast, RecSys data such as ZVUK exhibit more diverse and volatile trajectories (Figure 6, Appendix A.4), while the Shakespeare dataset, despite being textual, resembles banking data with largely flat drift patterns (Figure 7, Appendix A.4). Detailed analyses for individual datasets are provided in the Appendix A.4.

Motivation. These observations motivate a prevent-level *staticity index* that can be computed *before* model training to guide the choice of modeling strategy. Unlike the single-anchor variant (first window vs. all others), we adopt a more robust, multi-anchor formulation.

Staticity index. Fix a window length W and stride s . For each sequence u with per-window distributions $\{P_i^{(u)}\}_{i=1}^{I_u}$, choose anchors \mathcal{R}_u (uniformly at random, $R = 3$). The per-sequence score is the average shape-similarity

$$S^{(u)} = \frac{1}{RI_u} \sum_{r \in \mathcal{R}_u} \sum_{i=1}^{I_u} \text{Shape}(P_r^{(u)}, P_i^{(u)}),$$

and the dataset-level index is Staticity = $\frac{1}{N} \sum_{u=1}^N S^{(u)}$.

Thus, the staticity index quantifies the temporal stability of a sequence’s multi-feature distribution: higher values (close to 1) reflect stronger staticity (quasi-stationarity), whereas values near zero indicate pronounced drift. Importantly, the conclusions derived from the computed staticity index (Table 2) align with those previously inferred from the qualitative analysis of the plots.

4.3 LOCAL PERMUTATION OF EVENTS

To assess the importance of temporal order, we apply a local permutation operator parameterized by a window radius w . For each position i , we construct a symmetric window $[i - w, i + w]$, and the event at position i is allowed to move only within this window. Specifically, we construct a square cost matrix filled with random values, mask out entries corresponding to positions outside the window (and all padding tokens), and compute the constrained permutation via the Hungarian algorithm (Kuhn, 1955). We use $w \in \{0, 1, 4, 16, -1\}$, where larger values correspond to stronger

270 disruption of local order. The case $w = -1$ removes the positional constraint completely and allows
 271 a global permutation of the sequence.
 272

273 Importantly, even when $w = -1$, we strictly prevent mixing between the historical part of the se-
 274 quence and the target part: the two segments are permuted independently. This ensures that the
 275 model never sees target tokens reintroduced into the history during shuffling. For each window size
 276 w , we train and evaluate the model under the corresponding level of local permutation, enabling us
 277 to study how different datasets respond to disrupted temporal structure.
 278

279 4.4 OTHER STATISTICS

280 We also report the exponential decay parameter λ , which quantifies how quickly category frequen-
 281 cies decline in each dataset. Specifically, λ is the rate parameter of an exponential distribution fitted
 282 to the empirical histogram of event categories. This fit provides a compact measure of distributional
 283 imbalance: larger λ values indicate a steeper decay and, consequently, a stronger dominance of the
 284 most frequent categories. Figure 1 illustrates the fitted exponential curves alongside the empirical
 285 histograms for several datasets. The corresponding λ values for all datasets are reported in Table 2.
 286

287 Additionally, we report the Consecutive Repeat Rate (CRR)—the average length of uninterrupted
 288 runs of identical tokens. Higher CRR indicates more repetition, which can inflate short-term predic-
 289 tion accuracy. CRR values are listed in Table 2.
 290

291

292 5 DISTRIBUTION FORECASTING METHODS

293 We study the task of forecasting a distribution of a sequence over some horizon N given its history.
 294 To this end, we consider several training objectives — autoregressive, target-based, matched, and
 295 our order-invariant formulation. For all experiments N is fixed as 32.
 296

297

298 5.1 AUTOREGRESSIVE LOSS

300 Let $x_{1:T}$ be a sequence with $x_t \in \{1, \dots, K\}$. The model parameterises conditional next-event
 301 probabilities $p_\theta(x_{t+1} \mid x_{1:t})$ given the preceding context $x_{1:t}$. The sequence log-likelihood fac-
 302 torises as: $\log p_\theta(x_{1:T}) = \sum_{t=1}^T \log p_\theta(x_{t+1} \mid x_{1:t})$.
 303

304

305 5.2 TARGET LOSS

306 In this setting the model predicts an entire block of L future events *in a single forward pass*, using
 307 a fixed prefix $x_{1:T}$ as context; no teacher forcing is applied inside the horizon. Let $\hat{p}_{T+1}, \dots, \hat{p}_{T+L}$
 308 be the categorical distributions produced for positions $T+1$ through $T+L$. The target loss is the
 309 sum of negative log-likelihoods for that block: $\mathcal{L}_{\text{target}}^{(L)} = \sum_{i=T+1}^{T+L} -\log \hat{p}_i(x_i \mid x_{1:T})$
 310

311 Unlike the autoregressive objective, every term is conditioned on *the same* prefix $x_{1:T}$; the model
 312 **GRU-Target** therefore learns to produce an entire horizon coherently without receiving the ground-
 313 truth events $x_{T+1:T+L-1}$ as intermediate inputs.
 314

315

316 5.3 MATCHED LOSS

317 When the temporal order of future events is weakly informative, forcing the model to predict both
 318 the *events* and their *exact positions* needlessly penalises near-correct outputs. The **GRU-Matched**
 319 model adapts the matching idea of Karpukhin & Savchenko (2024), aligning each target event with
 320 the nearest prediction within a tolerance window of size m , treated as a hyperparameter.
 321

322 Let a fixed prefix $x_{1:T}$ condition a one-shot block prediction $\hat{p}_{T+1:T+L}$; let $x_{T+1:T+L}$ be the cor-
 323 responding ground truth. With a permutation σ constrained by $|\sigma(i) - i| \leq m$, the matched loss is

$$\mathcal{L}_{\text{match}}^{(m)} = \min_{\sigma \in \mathcal{A}} \sum_{i=T+1}^{T+L} -\log \hat{p}_{\sigma(i)}(x_i \mid x_{1:T}).$$

324 At $m = 0$ it reduces to plain block cross-entropy; as m grows, the objective becomes progressively
 325 order-invariant. The minimisation is solved with the Hungarian algorithm on the cost matrix $\ell_{ij} =$
 326 $-\log \hat{p}_j(x_i \mid x_{1:T})$.
 327

328 5.4 ORDER-INVARIANT DISTRIBUTION PARAMETERIZATION 329

330 When the order of future events is not informative, it is sufficient to model only the *event type*
 331 *distribution* rather than their precise temporal arrangement. We therefore introduce the **GRU-Dist**
 332 model, which represents each sequence as a *bag of events* and is trained to match the empirical
 333 distribution.

334 Let $H_t = \{x_1, \dots, x_t\}$ be the multiset of events observed so far. A neural encoder f_θ maps H_t
 335 to logits, which are converted to probabilities $\pi_t = \text{softmax}(f_\theta(H_t)) \in \Delta^{K-1}$, where Δ^{K-1}
 336 is the probability simplex in \mathbb{R}^K . For a sequence of length L we form its empirical distribution
 337 $\hat{p}_k = \frac{1}{L} \sum_{t=1}^L \mathbf{1}\{x_t = k\}$, and minimize $\ell(\theta) = D_{\text{KL}}(\hat{p} \parallel \pi(\theta))$.
 338

339 Unlike autoregressive objectives that require $L \times K$ logits per sequence, our order-invariant head
 340 outputs only a single K -dimensional vector. This reduces both computational and memory costs by
 341 a factor of L , while remaining well suited for datasets where event order carries little information.
 342

343 6 EVALUATION

344 For each configuration *Dataset* \times *Method* \times *LocalShuffle* we perform an extensive hyperparameter
 345 optimization of 100 trials, technical details are given in Appendix A.1.
 346

347 6.1 BASELINES

349 **We consider four simple baselines.** (1) **Ground Truth** uses the original sequences as a sanity
 350 check and reference point for metrics such as Cardinality. **Repeat** extends a sequence by copying
 351 its most recent observations into the forecast horizon of the length N . **Mode** outputs the users most
 352 frequent category for all N , illustrating the tendency of autoregressive models to collapse into trivial
 353 mode repetition—a behavior that may be overestimated by order-dependent metrics (e.g., Accuracy,
 354 Levenshtein distance). Finally, **HistSampler** generates sequences by sampling from the empirical
 355 histogram of past users sequence, thereby preserving marginal category frequencies while discarding
 356 temporal dependencies.

357 6.2 NEURAL BACKBONES

359 We evaluate two neural backbone architectures for sequence modeling:
 360

- 361 • **GRU:** The standard Gated Recurrent Unit (GRU) Cho et al. (2014) excels in capturing
 362 local dependencies and stationary patterns in short to moderately long time series.
- 363 • **GPT:** GPT-2 Radford et al. (2019), a causal Transformer-based model capable of modeling
 364 long-range dependencies, crucial for sequences with complex contextual interactions and
 365 implicit event relationships.

367 6.3 MULTI-TOKEN PREDICTION VIA SAMPLING

369 **Sampling in the order-sensitive models:** Autoregressive decoding with greedy argmax often col-
 370 lapses to the modal category. A simple remedy is to *sample* from the predictive categorical distri-
 371 bution instead of always taking the maximum, which reduces *mode collapse* and improves order-
 372 invariant metrics. For autoregressive and block-prediction models this sampling is straightforward,
 373 as logits at each step define the distribution, in our order-invariant method the distribution itself
 374 is parameterized directly, making sampling the natural decoding mechanism. We did not analyze
 375 more sophisticated sampling approaches such as beam search and our preliminary experiments with
 376 temperature sampling did not provide stable improvement across datasets, so we do not use them.
 377

Sampling in the order-invariant model: Given a predicted categorical distribution
 $\pi = (\pi_1, \dots, \pi_K)$ and a target length L , we first compute the expected fractional counts $\hat{n}_k = L\pi_k$.

378 Table 3: Next $N = 32$ tokens forecasting. *Matched-F1 (micro)* for all datasets and methods includ-
 379 ing baselines. \dagger denotes sampled version of method.

Method	MBD	Age	AB	Retail	ShS	Taobao	MM	Zvuk
GT	1.000	1.000	1.000	1.000	1.000	0.926	1.000	1.000
Mode	0.520	0.331	0.380	0.219	0.158	0.117	0.156	0.113
Repeat	0.830	0.680	0.700	0.661	0.587	0.257	0.318	0.274
HistSampler	0.804	0.632	0.680	0.640	0.533	0.197	0.244	0.226
GRU	0.528	0.477	0.375	0.207	0.596	0.222	0.250	0.148
GRU \dagger	0.771	0.628	0.641	0.609	0.596	0.146	0.171	0.126
GPT	0.524	0.476	0.373	0.212	0.594	0.223	0.250	0.192
GPT \dagger	0.776	0.627	0.629	0.611	0.603	0.151	0.188	0.174
GRU-Target	0.541	0.370	0.403	0.398	0.299	0.196	0.267	0.143
GRU-Target \dagger	0.808	0.633	0.670	0.641	0.572	0.154	0.201	0.140
GRU-Matched	0.847	0.704	0.676	0.708	0.688	0.203	0.272	0.202
GRU-Matched \dagger	0.827	0.653	0.647	0.667	0.634	0.155	0.203	0.134
GRU-Dist	0.856	0.725	0.736	0.719	0.705	0.178	0.247	0.239

398 Since these values are not integers, we obtain discrete category counts (n_1, \dots, n_K) using Hamil-
 399 ton’s method Balinski & Young (2010), $(n_1, \dots, n_K) = \text{Hamilton}(\hat{n}_1, \dots, \hat{n}_K)$, $\sum_{k=1}^K n_k = L$.
 400 This method distributes L discrete slots among categories in proportion to their predicted probabili-
 401 ties π_k and ensures that the total count equals L .

403 6.4 METRICS

405 Many classical sequence metrics (e.g., *Accuracy*, *Levenshtein distance*, *F1-score*) are defined with
 406 respect to a fixed token order and therefore penalize any permutation of events, even when such
 407 reordering is irrelevant for the problem at hand. To overcome this limitation, we introduce an order-
 408 invariant *Matched-F1* score, which treats sequences as *bags of events*.

409 To avoid order dependence we redefine true-positive, false-positive and false-negative terms. Let g_k
 410 and \hat{g}_k denote the ground-truth and predicted multiplicities of class k in the window. We set

$$412 (TP_k, FP_k, FN_k) = (\min(g_k, \hat{g}_k), \max(0, \hat{g}_k - g_k), \max(0, g_k - \hat{g}_k)).$$

414 Based on this definitions, we compute our *Matched-F1* with **micro**- and **macro**-averaging, anal-
 415 ogous to the conventional *F1-score* formulation. Detailed definition of this metric placed in Ap-
 416 pendix A.6.1

417 To assess diversity, we use **Cardinality** (see Appendix A.6.2), which measures the number of dis-
 418 tinct categories generated by the model. Low values signal *mode collapse*, while values close to the
 419 ground-truth indicate faithful event variety.

420 For completeness, we also report **Levenshtein distance**, an order-sensitive metric that, although less
 421 relevant to our setting, provides a complementary reference for order preserving methods (Table 4).

423 7 RESULTS

425 **Dataset-level statistics.** The staticity index serve as useful diagnostics for anticipating whether
 426 sequence order is relevant. Results are presented in Table 2. In banking datasets, a single modal
 427 category dominates—accounting for more than 50% of all events—leading to high values of both
 428 λ and the staticity index. This dominance is also associated with a pronounced performance drop
 429 under local permutations, suggesting limited reliance on sequential order.

431 **Local permutation experiments** (see Section 4.3) further corroborate these findings; results are
 432 shown in Figure 3. Shakespeare and Zvuk exhibit sharp performance degradation when sequences

432 Table 4: Next $N = 32$ tokens forecasting. *Levenshtein* for all datasets and methods including base-
 433 lines. \dagger denotes sampled version of method.

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are shuffled, indicating strong local sequential structure. In contrast, most banking datasets show little to no degradation, reflecting the irrelevance of event order. This trend is especially evident in Figure 2, which illustrates minimal perplexity degradation under shuffling for these datasets.

Matched-F1 performance. Order-invariant methods achieve the best overall performance on most datasets, significantly outperforming order-sensitive approaches (Table 3). *GRU-Dist* consistently outperforms *GRU-Matched*. Exceptions are Taobao and Megamarket, where *GRU-Dist* underperforms. These datasets exhibit a high Consecutive Repeat Rate (CRR, Table 2), and other models exploit this by repeating recent categories. *GRU-Dist*, by design, cannot leverage such local repetition. All learning methods struggle on Taobao, Megamarket, and Zvuk due to their very low exponential decay parameter λ and extremely high cardinality (Table 2). Here, *Repeat* baseline performs best. Sampling improves Matched-F1 for most order-sensitive models by alleviating *mode collapse*, but not on Taobao, Megamarket, and Zvuk (again due to low λ and high cardinality). Even with sampling, they remain inferior to order-invariant models.

Levenshtein performance. The *Mode* baseline is strong compared to order-sensitive methods in MBD, Age, AlphaBattle and Retail, highlighting the difficulty of modeling precise order (Table 4). As expected, order-agnostic models perform worse, since they impose no ordering constraints. Surprisingly, NTP models remain competitive on Taobao and Megamarket despite severe *mode collapse* (cardinality = 1): their local-mode predictions adapt better to fast distributional shifts than the static global mode, consistent with their low Staticity index S (Table 2).

8 CONCLUSION

Our study demonstrates that model performance in event-sequence forecasting depends strongly on dataset properties and on whether order-invariant or order-sensitive evaluation is appropriate.

When temporal order is largely irrelevant, order-sensitive methods suffer from *mode collapse*, performing similarly to the *Mode* baseline. In this case, order-invariant metrics are more appropriate. Under these metrics, *GRU-Dist* is generally the best, except when the category distribution is highly skewed (low λ) or the repetition is high (high CRR), where the *Repeat* baseline dominates.

When temporal structure is strong (low Staticity index S , high Consecutive Repeat Rate (CRR), significant Perplexity increase under local permutation), order-sensitive metrics become more appropriate, and autoregressive models are preferable, often outperforming other baselines even in the presence of severe *mode collapse*.

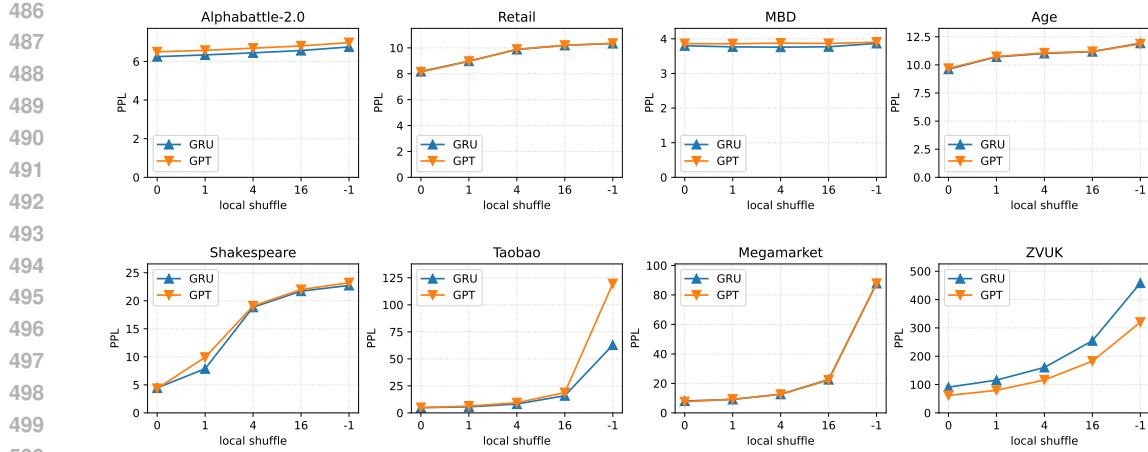
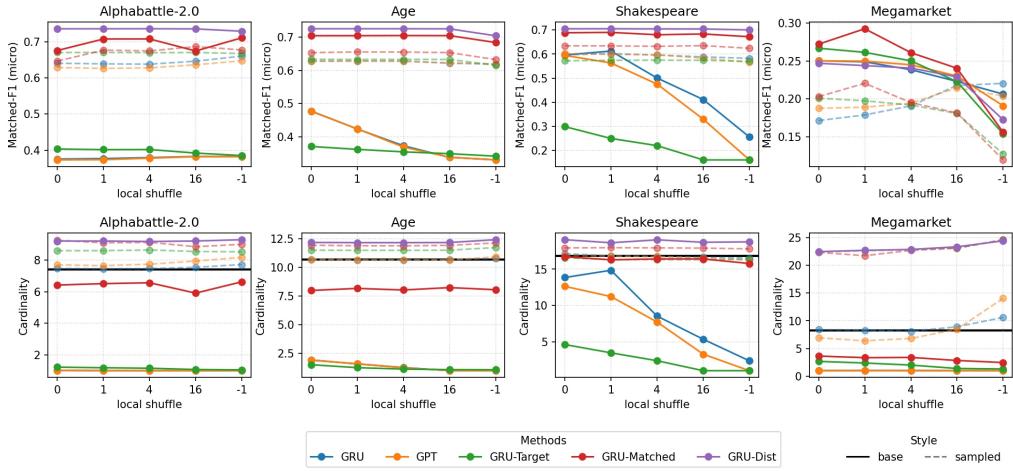
Figure 2: Next $N = 32$ tokens forecasting. Perplexity results.

Figure 3: Effect of Local Event Shuffling on Model Performance. We report Matched-F1 score and Cardinality for four datasets. Results for other datasets and metrics can be found in Appendix A.8

Cardinality also proves to be a useful diagnostic of *mode collapse*: in datasets like Shakespeare, shuffling removes structural cues and autoregressive models degenerate to the modal category. More broadly, when no meaningful local ordering exists, models tend to collapse to the mode (Figure 3).

Taken together, these results highlight the value of simple dataset-level diagnostics for anticipating model behavior, and demonstrate the advantages of order-invariant objectives in domains such as retail and banking, where event presence matters more than sequence order.

Indeed, it is worth noting that the proposed *GRU-Dist* method can be extended from single-category forecasting to multi-feature prediction through cascade modeling.

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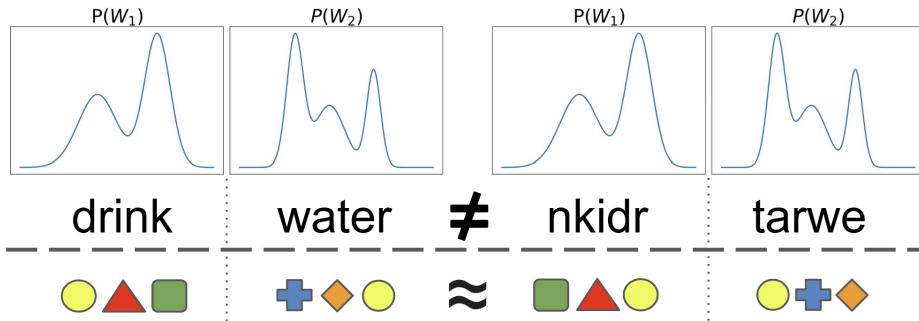
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648 **A APPENDIX**
649650 **A.1 HPO DETAILS**
651652 For hyperparameter optimization (HPO), we use Optuna (Akiba et al., 2019) with the Tree-structured
653 Parzen Estimator (TPE) sampler. For each model–dataset pair, we allocate an HPO budget of 100
654 training runs, capping the total computational cost at 18 NVIDIA A100 GPU-days. We reserve 15%
655 of the training set as a validation subset for early stopping and hyperparameter selection. The best-
656 performing hyperparameters are then used to train the final model for evaluation and all subsequent
657 study experiments.658 **A.2 LOCAL GLOBAL TEMPORAL INVARIANCE**
659660 In Figure 4 we illustrate local / global invariance.
661662
663 Figure 4: Example how order importance differs in different types of data. Even though in both cases
664 horizon distribution doesn't change, event sequence still make sense after permut inside intervals.
665666 **A.3 DATASETS DESCRIPTION AND PREPROCESSING**
667668 **MBD**¹ is a multimodal banking dataset introduced in Mollaev et al. (2024). The dataset contains
669 an industrial-scale number of sequences, with data from more than 1.5 million clients in 2 year
670 period. Each client corresponds to a sequence of events. This multi-modal dataset includes card
671 transactions, geo-position events, and embeddings of dialogs with technical support. For our analysis,
672 we use only card transactions. We use a temporal train–test split: transactions from the first year
673 form the training set, and those from the second year form the test set.674 **Age** dataset² consists of 44M anonymized credit card transactions representing 50K individuals.
675 The target is to predict the age group of a cardholder that made the transactions. Each transaction
676 includes the date, type, and amount being charged. The dataset was first introduced in scientific
677 literature in work Babaev et al. (2022). We perform a user-based split: 80% of sequences are
678 assigned to the training set, and the remaining 20% of sequences are held out for testing.679 **Retail** dataset³ comprises 45.8M retail purchases from 400K clients, with the aim of predicting
680 a client's age group based on their purchase history. Each purchase record includes details such
681 as time, item category, the cost, and loyalty program points received. The age group information
682 is available for all clients, and the distribution of these groups is balanced across the dataset. The
683 dataset was first introduced in scientific literature in work Babaev et al. (2022). We perform a user-
684 based split: 80% of sequences are assigned to the training set, and the remaining 20% of sequences
685 are held out for testing.686
687 ¹<https://huggingface.co/datasets/ai-lab/MBD>
688689 ²<https://ods.ai/competitions/sberbank-sirius-lesson>
690691 ³<https://ods.ai/competitions/x5-retailhero-uplift-modeling>
692

702 **Alphabattle-2.0** dataset⁴ The AlfaBattle 2.0 dataset contains bank customers’ transaction records
 703 over two years, with the goal of predicting loan default based on behavioral history. Each record in-
 704 cludes 18 features (3 numeric, 15 categorical) per transaction. We use the official test split provided
 705 by the dataset creators.

707 **Shakespeare** Dataset consists of character-level text extracted from Shakespeare’s works, prepro-
 708 cessed into individual speech segments. Each speech is tokenized using a vocabulary of unique
 709 characters mapped to integer codes based on frequency. The final dataset is split into train and test
 710 sets (80/20). The dataset is designed for character-level language modeling and was selected due to
 711 it obvious temporal importance.

712 **Zvuk** dataset⁵ is introduced in 2024 and contains 244.7M music listening events grouped into
 713 12.6M sessions from 382K users, recorded during the same five-month period (January–May 2023).
 714 In total, it spans 1.5M unique tracks. Each record includes a user ID, session ID, track ID, times-
 715 tamp, and play duration (considering only plays covering at least 30% of track length). The dataset
 716 is tailored to music consumption, excluding podcasts and audiobooks, and enables evaluation of rec-
 717 ommendation models in domains with stronger sequential dynamics. We use a temporal train–test
 718 split: transactions from the first two months form the training set, and other two month form the test
 719 set.

721 **MegaMarket** dataset⁶ is introduced in 2024 and comprises 196.6M user interactions collected
 722 over a five-month period (January–May 2023). It covers 2.7M users, 3.56M items, and 10,001
 723 product categories, with events including views, favorites, cart additions, and purchases. Each record
 724 contains a user ID, item ID, event type, category ID, timestamp, and normalized price. The dataset
 725 represents large-scale e-commerce behavior and is intended for sequential recommendation tasks.
 726 This dataset follows the same temporal train/test split as Zvuk.

727 **Taobao** ⁷ The dataset comprises user behaviors from Taobao, including clicks, purchases, adding
 728 items to the shopping cart, and favoriting items. These events were collected between November 18
 729 and December 15. The training set encompasses data from November 18 to December 1, while the
 730 test set includes clicks from December 2 to December 15.

732 A.4 STATICITY INDEX PLOTS FOR KEY DATASETS

734 For each dataset, we compute drift trajectories for all sequence and cluster them into a small number
 735 of groups with internally consistent dynamics (Figure 5–7). Across banking datasets (MBD, Retail,
 736 Age, Alphabattle) the dominant clusters are static, as exemplified for **MBD** (Figure 5c), these clus-
 737 ters exhibit negligible temporal drift. For such sequences, learning the user’s category distribution
 738 suffices to forecast the next block of events. Trajectories with pronounced drift are rare. In MBD
 739 specifically, such sequences are observed in fewer than 6% of users (Figure 5b).

740 In contrast to banking datasets, recommender–system data exhibit much greater variability. In
 741 **ZVUK** (Figure 6), two characteristic regimes dominate: one cluster shows a sharp initial drop from
 742 the baseline followed by persistent high-variance fluctuations, while another appears quasi-static yet
 743 remains noisy around its trend. Such patterns reflect the broader nature of recommender logs: users
 744 interact with a large and diverse sets of items, and their behavior shifts more frequently than in retail
 745 domains where event types are limited and highly regular. And as a consequence, their later-window
 746 distributions are more clearly separated from the first-window distribution.

747 The outlier in this collection is the **Shakespeare** text dataset (Figure 7). Although it is non-
 748 transactional, its dynamics resemble banking data more than recommender logs: drift trajectories
 749 are mostly flat and volatility remains low. At the same time, weak periodic or gradual shifts are
 750 observable, indicating that the sequences are not fully static but display a modest degree of temporal
 751 variation.

752 ⁴<https://www.kaggle.com/datasets/mrmorj/alfabattle-20>

753 ⁵<https://www.kaggle.com/datasets/alexxl/zvuk-dataset>

754 ⁶<https://www.kaggle.com/datasets/alexxl/megamarket?select=megamarket.parquet>

755 ⁷<https://tianchi.aliyun.com/dataset/46>

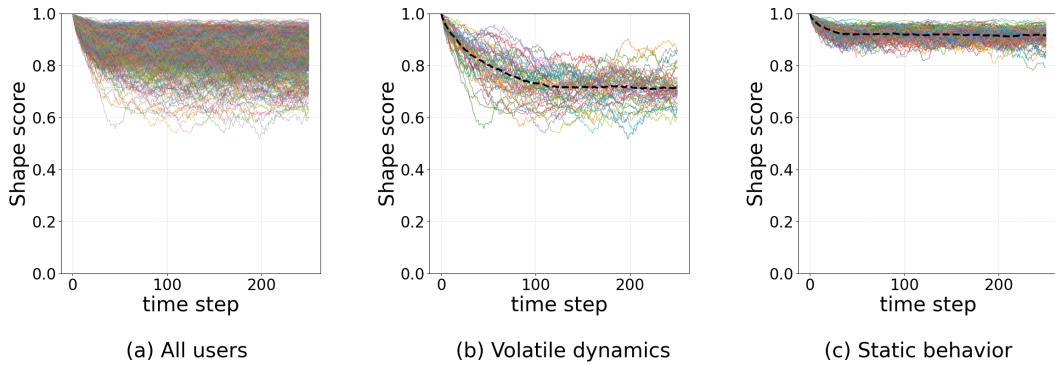


Figure 5: Shape score drift for MBD dataset

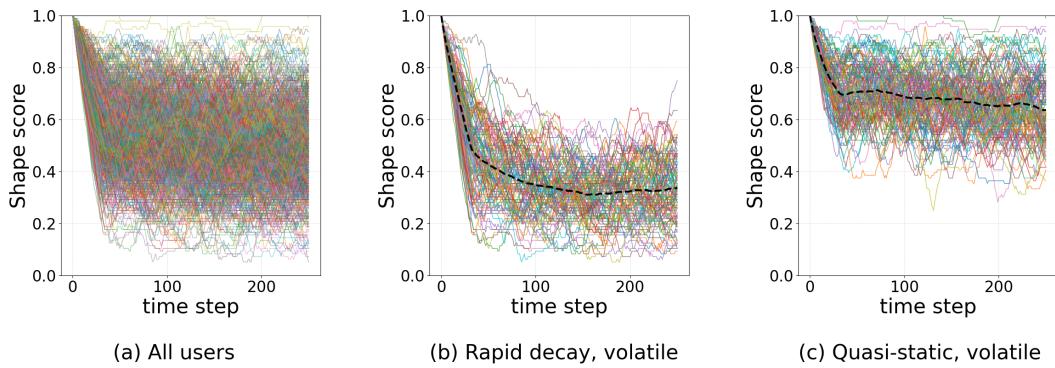


Figure 6: Shape score drift for ZVUK dataset

787 A.5 FEATURES IMPACT IN CATEGORY FORECASTING QUALITY

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789 We investigated whether predicting a target feature benefits more from incorporating the full feature
790 vector or from relying exclusively on its own historical values.
791

792 On the MBD dataset, experiments in the *All-to-One* and *One-to-One* modes reveal that the autoregressive
793 model’s performance degrades when exposed to complete with the complete feature vector.
794 The additional inputs act as noise, impeding the model’s ability to reproduce the mode of the target
795 distribution. In the *One-to-One* mode—where the model sees only the history of the target
796 feature—it easily learns the mode and reports a formal increase in accuracy; however, this gain is
797 illusory, as the generated sequences become overly uniform and lack realism 5.
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Table 5: Effect of training with all tokens vs. event type only (*Matched-F1 micro*).

Dataset	Change (%)
MBD	+2.85
AGE	-24.94
MM	+13.66

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807 By contrast, on datasets with a strong sequential structure, such as *Megamarket*, the opposite pattern
808 emerges. The autoregressive mechanism leverages ordering information and, when augmented with
809 additional features, predicts beyond mere modal values, resulting in a significant improvement in
810 performance metrics.

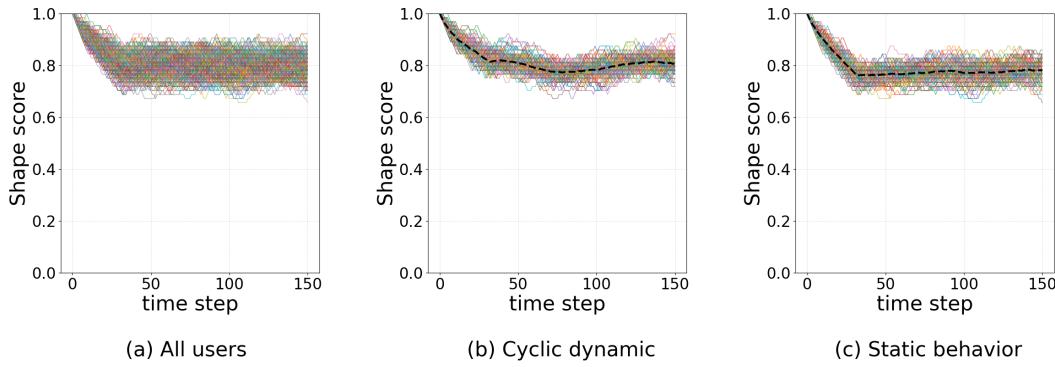


Figure 7: Shape score drift for Shakespeare dataset

A.6 METRICS

A.6.1 MATCHED-F1 MICRO

Precision and recall.

$$\text{Prec}_k = \frac{\text{TP}_k}{\text{TP}_k + \text{FP}_k}, \quad \text{Rec}_k = \frac{\text{TP}_k}{\text{TP}_k + \text{FN}_k}.$$

Macro averaging.

$$F1_{\text{macro}} = \frac{1}{K} \sum_{k=1}^K \frac{2 \text{Prec}_k \text{Rec}_k}{\text{Prec}_k + \text{Rec}_k}.$$

Each class contributes equally; the score is sensitive to rare categories.

Micro averaging. Aggregating counts over classes,

$$\text{TP} = \sum_k \text{TP}_k, \quad \text{FP} = \sum_k \text{FP}_k, \quad \text{FN} = \sum_k \text{FN}_k, \quad (2)$$

$$F1_{\text{micro}} = \frac{2 \text{TP}}{2 \text{TP} + \text{FP} + \text{FN}}. \quad (3)$$

This variant weights categories by frequency and reflects overall throughput.

A.6.2 CARDINALITY METRIC.

Let $G_i = (x_{t+1}^{(i)}, \dots, x_{t+L}^{(i)})$ denote the L -step segment generated for sequence i and $\mathcal{C}(G_i) = \{x \in G_i\}$ the set of *distinct* categories appearing in that segment. We define the per-sequence cardinality as

$$C_i = |\mathcal{C}(G_i)|.$$

The dataset-level score is the average

$$\text{Cardinality} = \frac{1}{N} \sum_{i=1}^N C_i,$$

where N is the number of sequences under evaluation. An *overall* variant first concatenates all generated segments, $\tilde{G} = \bigcup_i G_i$, and reports $C_{\text{overall}} = |\mathcal{C}(\tilde{G})|$.

Purpose. Cardinality captures the *category diversity* produced by a model: low values signal *mode collapse*, whereas values close to the ground-truth cardinality indicate faithful reproduction of event variety. We compute the metric for both generated (C_{gen}) and reference (C_{orig}) sequences, allowing direct comparison of a model's diversity against empirical data.

864 Table 6: Comparison of autoregressive baselines under the original one-to-one setup and the all-to-
 865 all variant. Metrics are computed only on the target event category.
 866

Metric	MBD	AGE	AB	Retail	ShS	Taobao	MM	Zvuk
Matched F1 (one-to-one)	0.528	0.476	0.375	0.208	0.596	0.222	0.250	0.148
Matched F1 (all-to-all)	0.440	0.474	0.375	0.006	0.612	0.176	0.026	0.004
Levenshtein (one-to-one)	0.520	0.390	0.374	0.194	0.200	0.222	0.250	0.139
Levenshtein (all-to-all)	0.440	0.391	0.373	0.005	0.203	0.176	0.026	0.004

872
 873 A.7 EFFECT OF INPUT SPECIFICATION FOR AUTOREGRESSIVE BASELINES
 874

875 In the main experiments, multi-token prediction models (GRU-Dist, GRU-Matched, and GRU-Tar-
 876 get) are trained in an all-to-one setting, where the model predicts the entire future window of tar-
 877 get categories given the history with all features in datasets including timestamps. Autoregressive
 878 models, however, cannot be trained in an all-to-one formulation: their training objective requires
 879 predicting a single event at a time and conditioning each step on the previously generated outputs.
 880 Therefore, all autoregressive baselines are trained in the standard one-to-one setting. This architec-
 881 tural restriction leads to a mild asymmetry in the input setup.
 882

883 To verify that this asymmetry does not influence our conclusions, we conducted an additional ex-
 884 periment in which the autoregressive baselines were trained in an all-to-all setting. In this variant,
 885 the model is provided with all future features in the prediction window, while the evaluation metrics
 886 (Matched-F1 and Levenshtein) are computed only on the target category, keeping the evaluation
 887 protocol unchanged.

888 Table 6 reports the comparison between the original one-to-one setup and the all-to-all variant.
 889 Across datasets, the all-to-all formulation does not improve autoregressive models. Crucially, the
 890 relative ranking between autoregressive and multi-token prediction approaches remains unchanged.
 891 This confirms that the advantages of GRU-Dist and related models are robust to the choice of input
 892 formulation.
 893

894 A.8 ADDITIONAL RESULTS
 895

896 For completeness, we report all evaluation metrics across datasets. Levenshtein distance is included
 897 as an order-sensitive measure to quantify degradation under local shuffling (Figure 8), while the
 898 effect of shuffling on category diversity is illustrated by cardinality (Figure 9). The main text focuses
 899 on the order-invariant *Matched-F1 (micro)* (Figure 10), which we adopt as the primary evaluation
 900 metric throughout the study.
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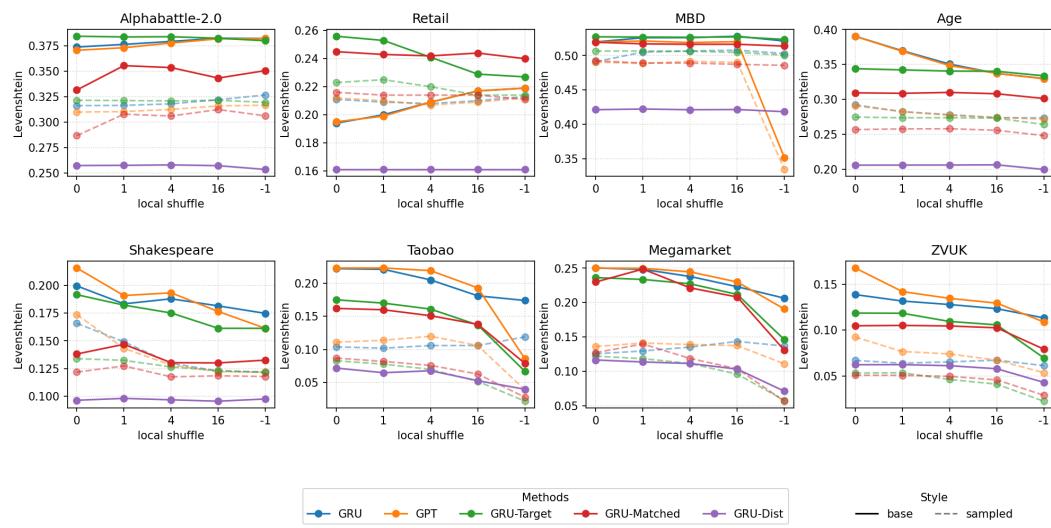


Figure 8: Levenshtein score on all datasets.

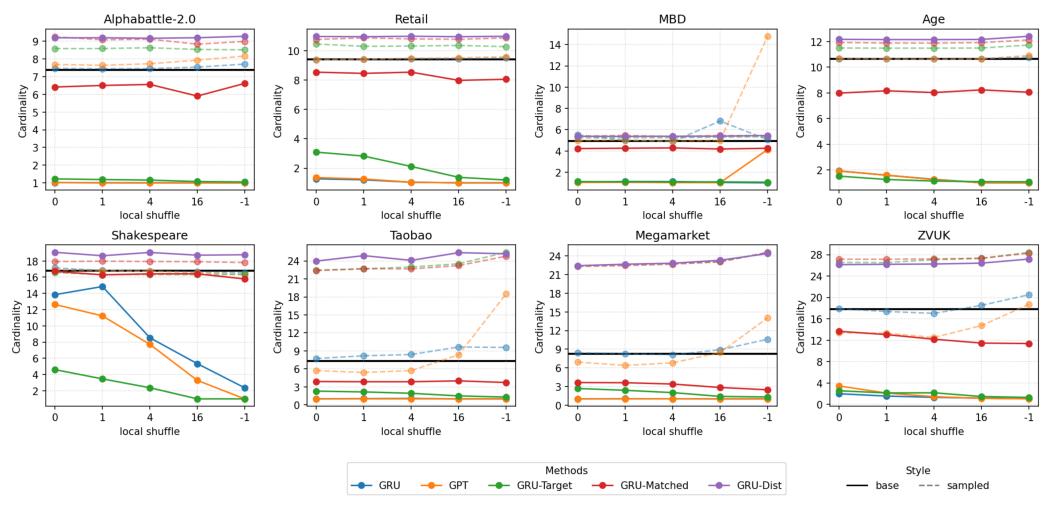
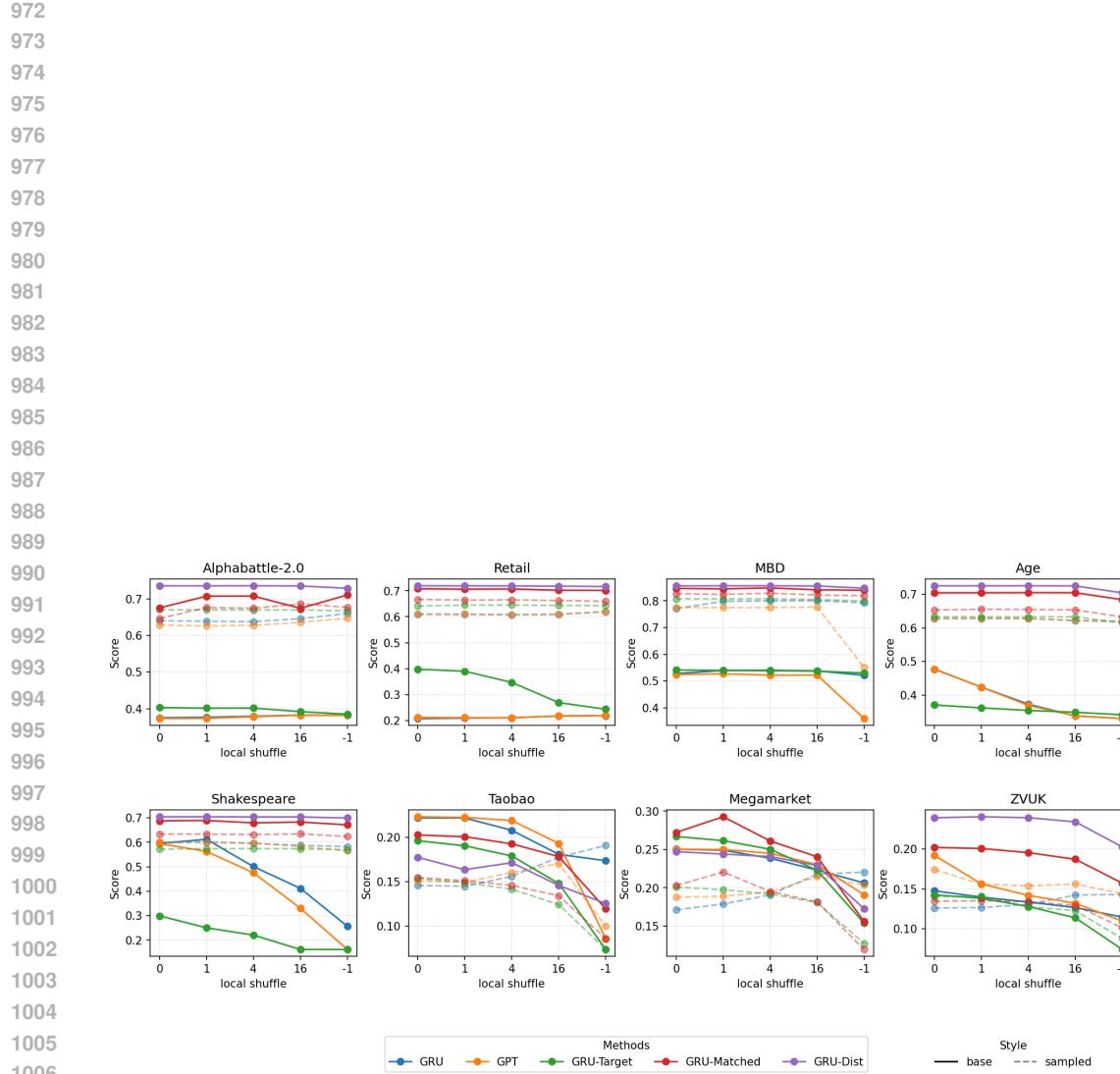


Figure 9: Effect of local shuffle on cardinality.

Figure 10: Next N tokens forecasting. *Matched-F1 (micro)* results.