SELECTIVE MIXUP HELPS WITH DISTRIBUTION SHIFTS, BUT NOT (ONLY) BECAUSE OF MIXUP

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Abstract

Context. Mixup is a highly successful technique to improve generalization of neural networks by augmenting the training data with combinations of random pairs. Selective mixup is a family of methods that apply mixup to specific pairs, e.g. only combining examples across classes or domains. These methods have claimed remarkable improvements on benchmarks with distribution shifts, but their mechanisms and limitations remain poorly understood.

Findings. We examine an overlooked aspect of selective mixup that explains its success in a completely new light. We find that the non-random selection of pairs affects the training distribution and improve generalization by means completely unrelated to the mixing. For example in binary classification, mixup across classes implicitly resamples the data for a uniform class distribution — a classical solution to label shift. We show empirically that this implicit resampling explains much of the improvements in prior work. Theoretically, these results rely on a "regression toward the mean", an accidental property that we identify in several datasets.

Takeaways. We have found a new equivalence between two successful methods: selective mixup and resampling. We identify limits of the former, confirm the effectiveness of the latter, and find better combinations of their respective benefits.

1 INTRODUCTION

Mixup and its variants are some of the few methods that improve generalization across tasks and modalities with no domain-specific information (Zhang et al., 2017). Standard mixup replaces training data with linear combinations of random pairs of examples, proving successful e.g. for image classification (Yun et al., 2019b), semantic segmentation (Islam et al., 2023), natural language processing (Verma et al., 2019), and speech processing (Meng et al., 2021).

This paper focuses on scenarios of distribution shift and variants of mixup that improve out-ofdistribution (OOD) generalization. We examine the family of methods that apply mixup on selected pairs of examples, which we refer to as *selective mixup* (Hwang et al., 2022; Li et al., 2023; Lu et al., 2022a; Palakkadavath et al., 2022; Tian et al., 2023; Xu et al., 2020; Yao et al., 2022b). Each method uses a predefined criterion,¹ for example combining examples across classes (Yao et al., 2022b) (Figure 1) or across domains (Xu et al., 2020; Li et al., 2023; Lu et al., 2022a). These simple heuristics have claimed remarkable improvements on benchmarks such as DomainBed (Gulrajani and Lopez-Paz, 2020), WILDS (Koh et al., 2021), and Wild-Time (Yao et al., 2022a).

Despite impressive empirical performance, the theoretical mechanisms of selective mixup remain obscure. For example, the selection criteria in Yao et al. (2022b) include the selection of pairs of the same class/different domains but also the exact opposite. This raises questions:

- 1. What makes each selection criterion suitable to any specific dataset?
- 2. Are there multiple mechanisms responsible for the improvements with selective mixup?

This paper presents surprising answers, highlighting an overlooked side effect of selective mixup. **The non-random selection of pairs implicitly biases the training distribution and improve generalization by means completely unrelated to the mixing**. We observe empirically that simply forming mini-batches with all instances of the selected pairs (without mixing them) often produces the same improvements as mixing them. This critical ablation was absent from prior studies.

¹We focus on the basic implementation (Yao et al., 2022b) without modifications to the learning objective.



Figure 1: Selective mixup is a family of methods that replace the training data with combined pairs of examples fulfilling a predefined criterion, e.g. pairs from different classes. An overlooked side effect is to modify the training distribution: here, sampling classes more uniformly. This is responsible for much of the observed improvements in OOD generalization.

We also analyze theoretically the resampling induced by different selection criteria. We find that conditioning on a "different attribute" (e.g. combining examples across classes or domains) brings the training distribution of this attribute closer to a uniform one. Consequently, the imbalances in the data often "regress toward the mean" with selective mixup. We verify empirically that several datasets do indeed shift toward a uniform class distribution in their test split (see Figure 10). We also find remarkable correlation between improvements in performance and the reduction in divergence of training/test distributions due to selective mixup. **This also predicts a new failure mode of selective mixup when the above property does not hold** (see Appendix C).

Our contributions are summarized as follows.

- We point out an overlooked resampling effect when applying selective mixup (Section 3).
- We show theoretically that certain selection criteria induce a bias in the distribution of features and/or classes equivalent to a "regression toward the mean" (Theorem 3.1). In binary classification for example, selecting pairs across classes is equivalent to sampling uniformly over classes, the standard approach to address label shift and imbalanced data.
- We verify empirically that multiple datasets indeed contain a regression toward a uniform class distribution across training and test splits (Section 4.6). We also find that improvements from selective mixup correlate with reductions in divergence of training/test distributions over labels and/or covariates. This strongly suggests that resampling is the main driver for these improvements.
- We compare many selection criteria and resampling baselines on five datasets. In all cases, improvements with selective mixup are partly or fully explained by resampling effects (Section 4).

The implications for future research are summarized as follows.

- We connect two areas of the literature by showing that selective mixup is sometimes equivalent to resampling, a classical strategy for distribution shifts (Garg et al., 2023; Idrissi et al., 2022). This hints at possible benefits from advanced methods for label shift and domain adaptation on benchmarks with distribution shifts.
- The resampling explains why different criteria in selective mixup benefit different datasets: they affect distributions of features and/or labels thus addressing covariate/label shift.
- This explanation highlights the risk of overfitting to the benchmarks: much of the improvements rely on the accidental "regression toward the mean" in the datasets examined.

2 BACKGROUND: MIXUP AND SELECTIVE MIXUP

Notations. We consider a classification model $f_{\theta} : \mathbb{R}^d \to [0, 1]^C$ of learned parameters θ . It maps an input vector $\boldsymbol{x} \in \mathbb{R}^d$ to a vector \boldsymbol{y} of scores over C classes. The training data is typically a set of labeled examples $\mathcal{D} = \{(\boldsymbol{x}_i, \boldsymbol{y}_i, d_i)\}_{i=1}^n$ where \boldsymbol{y}_i are one-hot vectors encoding ground-truth labels, and $d_i \in \mathbb{N}$ are optional discrete domain indices. Domain labels are available e.g. in datasets with different image styles (Li et al., 2017) or collected over different time periods (Koh et al., 2021).

Training with ERM. Standard empirical risk minimization (ERM) optimizes the model's parameters for $\min_{\theta} \mathcal{R}(f_{\theta}, \mathcal{D})$. The expected training risk for a chosen loss function \mathcal{L} is:

$$\mathcal{R}(f_{\boldsymbol{\theta}}, \mathcal{D}) = \mathbb{E}_{(\boldsymbol{x}, \boldsymbol{y}) \in \mathcal{D}} \mathcal{L}(f_{\boldsymbol{\theta}}(\boldsymbol{x}), \boldsymbol{y}).$$
(1)

An empirical estimate is obtained with an arithmetic mean over instances of the dataset \mathcal{D} .

Training with mixup. Standard mixup essentially replaces training examples with linear combinations of random pairs in both input and label space. We formalize it by redefining the training risk:

$$\mathcal{R}_{\text{mixup}}(f_{\boldsymbol{\theta}}, \mathcal{D}) = \mathbb{E}_{(\boldsymbol{x}, \boldsymbol{y}) \in \mathcal{D}} \mathcal{L}(f(c \, \boldsymbol{x} + (1-c) \widetilde{\boldsymbol{x}}, c \, \boldsymbol{y} + (1-c) \, \widetilde{\boldsymbol{y}}))$$
(2)

with mixing coefficients $c \sim \mathcal{B}(2,2)$ and paired examples $(\tilde{x}, \tilde{y}) \sim \mathcal{D}$. (3)

The expectation is approximated by sampling coefficients and pairs at every training iteration.

Selective mixup. While standard mixup combines random pairs, selective mixup only combines pairs that fulfill a predefined criterion. To select these pairs, the method starts with the original data \mathcal{D} , then for every $(x, y, d) \in \mathcal{D}$ it selects a $(\tilde{x}, \tilde{y}, \tilde{d}) \in \mathcal{D}$ such that they fulfill the criterion represented by the predicate Paired (\cdot, \cdot) . For example, the criterion *same class, different domain* ("intra-label LISA" in Yao et al. (2022b)) is implemented as:

 $\operatorname{Paired}((\boldsymbol{x}_i, \boldsymbol{y}_i, d_i), (\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i, \widetilde{d}_i)) = true \text{ iff } (\widetilde{\boldsymbol{y}} = \boldsymbol{y}) \land (\widetilde{d} \neq d) \text{ (same class, diff. domain)}$ (4a)

Other examples:

$$\operatorname{Paired}((\boldsymbol{x}_i, \boldsymbol{y}_i, d_i), (\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i, d_i)) = true \text{ iff } (\widetilde{\boldsymbol{y}} \neq \boldsymbol{y}) \qquad (different \ class) \tag{4b}$$

 $\operatorname{Paired}((\boldsymbol{x}_i, \boldsymbol{y}_i, d_i), (\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i, \widetilde{d}_i)) = true \text{ iff } (\widetilde{d} = d) \qquad (same \ domain) \tag{4c}$

3 SELECTIVE MIXUP MODIFIES THE TRAINING DISTRIBUTION

The new claims of this paper comprise two parts.

- 1. Estimating the training risk with selective mixup (Eq. 2) uses a different sampling of examples from \mathcal{D} than ERM (Eq. 1). We demonstrate this theoretically in this section.
- 2. We hypothesize that this different sampling of training examples influences the generalization properties of the learned model, regardless of the mixing operation. We verify this empirically in Section 4 using ablations of selective mixup that omit the mixing operation a critical baseline absent from prior studies.

Training distribution. This distribution refers to the examples sampled from \mathcal{D} to estimate the training risk (Eq. 1 or 2) — whether these are then mixed or not. The following discussion focuses on distributions over classes (y) but analogous arguments apply to covariates (x) and domains (d).

With ERM, the training distribution equals the dataset distribution because the expectation in Eq. (1) is over uniform samples of \mathcal{D} . We obtain an empirical estimate by averaging all one-hot labels, giving the vector of discrete probabilities $\mathbf{p}_{\mathbf{Y}}(\mathcal{D}) = \bigoplus_{(\boldsymbol{x},\boldsymbol{y})\in\mathcal{D}} \boldsymbol{y} / |\mathcal{D}|$ where \oplus is the element-wise sum.

With selective mixup, evaluating the risk (Eq. 2) requires pairs of samples. The first element of a pair is sampled uniformly, yielding the same $\mathbf{p}_{\mathbf{Y}}(\mathcal{D})$ as ERM. The second element is selected as described above, using the first element and one chosen predicate $\operatorname{Paired}(\cdot, \cdot)$ e.g. from (4a–4c). For our analysis, we denote these "second elements" of the pairs as the virtual data:

$$\widetilde{\mathcal{D}} = \left\{ (\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i, \widetilde{d}_i) \sim \mathcal{D} : \text{Paired} \left((\boldsymbol{x}_i, \boldsymbol{y}_i, d_i), (\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i, \widetilde{d}_i) \right) = true, \quad \forall i = 1, \dots, |\mathcal{D}| \right\}.$$
(5)

We can now analyze the overall training distribution of selective mixup. An empirical estimate is obtained by combining the distributions resulting from the two elements of the pairs, which gives the vector $\mathbf{p}_{\mathbf{Y}}(\mathcal{D} \cup \widetilde{\mathcal{D}}) = (\mathbf{p}_{\mathbf{Y}}(\mathcal{D}) \oplus \mathbf{p}_{\mathbf{Y}}(\widetilde{\mathcal{D}}))/2$.

Regression toward the mean. With the criterion same class, it is obvious that $\mathbf{p}_{\mathbf{Y}}(\widetilde{\mathcal{D}}) = \mathbf{p}_{\mathbf{Y}}(\mathcal{D})$. Therefore these variants of selective mixup are not concerned with resampling effects.² In contrast, the criteria different class or different domain do bias the sampling. In the case of binary classification, we have $\mathbf{p}_{\mathbf{Y}}(\widetilde{\mathcal{D}}) = 1 - \mathbf{p}_{\mathbf{Y}}(\mathcal{D})$ and therefore $\mathbf{p}_{\mathbf{Y}}(\mathcal{D} \cup \widetilde{\mathcal{D}})$ is uniform. This means that selective mixup with the different class criterion has the side effect of balancing the training distribution of classes, a classical mitigation of class imbalance (Japkowicz, 2000; Kubat et al., 1997). For multiple classes, we have a more general result.

Theorem 3.1. Given a dataset $\mathcal{D} = \{(\boldsymbol{x}_i, \boldsymbol{y}_i)\}_i$ and paired data $\widetilde{\mathcal{D}}$ sampled according to the "different class" criterion, i.e. $\widetilde{\mathcal{D}} = \{(\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i) \sim \mathcal{D} \text{ s.t. } \widetilde{\boldsymbol{y}}_i \neq \boldsymbol{y}_i\}$, then the distribution of classes in $\mathcal{D} \cup \widetilde{\mathcal{D}}$ is

²The absence of resampling effects holds for *same class* and *same domain* alone, but not in conjunction with other criteria. See e.g. the differences between *same domain/diff. class* and *any domain/diff. class* in Figure 3.

more uniform than in \mathcal{D} . Formally, the entropy $\mathbb{H}(\mathbf{p}_{Y}(\mathcal{D})) \leq \mathbb{H}(\mathbf{p}_{Y}(\mathcal{D} \cup \widetilde{\mathcal{D}}))$. *Proof:* see Appendix D.

Theorem 3.1 readily extends in two ways. First, the same effect also results from the *different domain* criterion: if each domain contains a different class distribution, the resampling from this criterion averages them out, yielding a more uniform aggregated training distribution. Second, this averaging applies not only to class labels (y) but also covariates (x). An analysis using distributions is ill-suited but the mechanism similarly affects the sampling of covariates when training with selective mixup.

When does one benefit from the resampling (regardless of mixup)? The above results mean that selective mixup can implicitly reduce imbalances (a.k.a. biases) in the training data. When these are not spurious and also exist in the test data, the effect on predictive performance could be detrimental.

We expect benefits (verified in Section 4) on datasets with distribution shifts. By definition, their training/test splits contain different imbalances. Softening imbalances in the training data is then likely to bring the training and test distributions closer, in particular with extreme shifts such as the complete reversal of a spurious correlation (e.g. *waterbirds* dataset, see Section 4.1).

We also expect benefits on worst-group metrics (e.g. *civilComments* dataset, see Section 4.4). The challenge in these datasets comes from the imbalance of class/domain combinations. Prior work has indeed shown that balancing is beneficial (Idrissi et al., 2022; Sagawa et al., 2019).

4 EXPERIMENTS

We performed a large number of experiments to understand the contribution of the different effects of selective mixup and other resampling baselines (complete results in Appendix B).

Datasets. We focus on five datasets that previously showed improvements with selective mixup. We selected them to cover a range of modalities (vision, NLP, tabular), settings (binary, multiclass), and types of shifts (covariate, label, and subpopulation shifts).

- Waterbirds (Sagawa et al., 2019) is a popular artificial dataset used to study distribution shifts. The task is to classify images of birds into two types. The image backgrounds are also of two types, and the correlation between birds and backgrounds is reversed across the training and test splits. The type of background in each image serves as its domain label.
- **CivilComments** (Koh et al., 2021) is a widely-used dataset of online text comments to be classified as toxic or not. Each example is labeled with a topical attribute (e.g. Christian, male, LGBT, etc.) that is spuriously associated with ground truth labels in the training data. These attributes serve as domain labels. The target metric is the worst-group accuracy where the groups correspond to all toxicity/attribute combinations.
- Wild-Time Yearbook (Yao et al., 2022a) contains yearbook portraits to be classified as male or female. It is part of the Wild-Time benchmark, which is a collection of real-world datasets captured over time. Each example belongs to a discrete time period that serves as its domain label. Distinct time periods are assigned to the training and OOD test splits (see Figure 10).
- Wild-Time arXiv (Yao et al., 2022a) contains titles of arXiv preprints. The task is to predict each paper's category among 172 classes. Time periods serve as domain labels.
- Wild-Time MIMIC-Readmission (Yao et al., 2022a) contains hospital records (sequences of codes representing diagnoses and treatments) to be classified into two classes. The positive class indicates the readmission of the patient at the hospital within 15 days. Time periods serve as domain labels.

Methods. We train standard architectures suited to each dataset with the methods below (details in Appendix A). We perform early stopping i.e. recording metrics for each run at the epoch of highest ID or worst-group validation performance (for *Wild-Time* and *waterbirds/civilComments* datasets respectively). We plot average metrics in bar charts over 9 different seeds with error bars representing \pm one standard deviation. **ERM** and **vanilla mixup** are the standard baselines. Baseline **resampling** uses training examples with equal probability from each class, domain, or combinations thereof as in Idrissi et al. (2022); Sagawa et al. (2019). **Selective mixup** (\blacksquare) includes all possible selection criteria based on classes and domains. We avoid ambiguous terminology from earlier works because of inconsistent usage (e.g. "intra-label LISA" means "different domain" in Koh et al. (2021) but not in Yao et al. (2022a)). **Selective sampling** (\blacksquare) is a novel ablation of selective mixup where the selected pairs are not mixed, but the instances are appended one after another in the mini-batch. Half

are dropped at random to keep the mini-batch size identical to the other methods. Therefore any difference between selective sampling and ERM is attributable only to resampling effects. We also include **novel combinations** (\blacksquare) of sampling and mixup.

4.1 **RESULTS ON THE** *waterbirds* DATASET

The target metric for this dataset is the worst-group accuracy, with groups defined as the four class/domain combinations. The two difficulties are (1) a class imbalance (77/23%) and (2) a correlation shift (spurious class/domain association reversed at test time). See discussion in Figure 2.



Figure 2: Main results on waterbirds.

We first observe that vanilla mixup is detrimental compared to ERM. Resampling with uniform class/domain combinations is hugely beneficial, for the reasons explained in Figure 3. The ranking of various criteria for selective sampling is similar whether with or without mixup. Most interestingly, the best criterion performs similarly, but no better than the best resampling.



The excellent performance of the best version of selective mixup is here entirely due to resampling. The efficacy of resampling on this dataset is not a new finding (Idrissi et al., 2022; Sagawa et al., 2019). What is new is its equivalence with the best variant of selective mixup. Figure 3 further supports this claim by comparing proportions of classes and domains sampled by each method.



Figure 3: The sampling ratios of each class/domain clearly explain the performance of the best methods (*waterbirds*).

4.2 **RESULTS ON THE** *yearbook* DATASET

The difficulty of this dataset comes from a slight class imbalance and the presence of covariate/label shift (see Figure 10). The test split contains several domains (time periods). The target metric is the worst-domain accuracy. Figure 4 shows that vanilla mixup is slightly detrimental compared to ERM. Resampling for uniform classes gives a clear improvement because of the class imbalance. With selective sampling (no mixup), the only criteria that improve over ERM contain "different class". This is expected because this criterion implicitly resamples for a uniform class distribution.

To investigate whether some of the improvements are due to resampling, we measure the divergence between training and test distributions of classes and covariates (details in Appendix A). Figure 5) shows first that there is a clear variation among different criteria (• blue dots) i.e. some bring the training/test distributions closer to one another. Second, there is a remarkable correlation between the test accuracy and the divergence, on both classes and covariates.³ This means that resampling effects do occur and also play a part in the best variants of selective mixup.

Finally, the improvements from simple resampling and the best variant of selective mixup suggest a new combination. We train a model with uniform class sampling and selective mixup using the

Resampling uniform combinations gives them all equal weights, just like the worst-group target metric. Selective mixup with same domain / diff. class also gives equal weights to the classes, while breaking the spurious pattern between groups and classes, unlike any other criterion.

³As expected, the correlation is reversed for the first two test domains in Figure 5 since they are even further from a uniform class distribution than the average of the training data, as seen in Figure 10.





Figure 5: Different selection criteria (•) modify the distribution of both covariates and labels (upper and lower rows). The resulting reductions in divergence between training and test distributions correlate remarkably well with test performance.³ This confirms the contribution of resampling to the overall performance of selective mixup.

"same class" criterion, and obtain performance superior to all existing results (last row in Figure 5). This confirms the **complementarity of the effects of resampling and within-class selective mixup**.

4.3 RESULTS ON THE arXiv DATASET

This dataset has difficulties similar to *yearbook* and also many more classes (172). Simple resampling for uniform classes is very bad (literally off the chart in Figure 6) because it overcorrects the imbalance (the test distribution being closer to the training than to a uniform one). Uniform *domains* is much better since its effect is similar but milder.

All variants of selective mixup (\blacksquare) perform very well, but they improve over ERM even without mixup (\blacksquare). And the selection criteria rank similarly with or without mixup, suggesting that parts of the improvements of selective mixup is due to the resampling. Given that vanilla mixup also clearly improves over ERM, the performance of **selective mixup is explained by cumulative effects of vanilla mixup and resampling effects**. This also suggests new combinations of methods (\blacksquare) among which we find one version marginally better than the best variant of selective mixup (last row).

4.4 RESULTS ON THE civilComments DATASET

This dataset mimics a subpopulation shift because the worst-group metric requires high accuracy on classes and domains under-represented in the training data. It also contains an implicit correlation shift because any class/domain association (e.g. "Christian" comments labeled as toxic more often than not) becomes spurious when evaluating individual class/domain combinations.

4.5 RESULTS ON THE MIMIC-Readmission DATASET

This dataset contains a class imbalance (about 78/22% in training data), label shift (the distribution being more balanced in the test split), and possibly covariate shift. It is unclear whether the task is causal or anticausal (labels causing the features) because the inputs contain both diagnoses and



Figure 6: Main results on arXiv.



To investigate the contribution of resampling, we measure the divergence between training/test class distributions and plot them against the test accuracy (Figure 7). We observe a strong correlation across methods. Mixup essentially offsets the performance by a constant factor. This suggests again the independence of the effects of mixup and resampling.



Figure 7: Divergence of tr./test class distributions vs. test accuracy.

The resampling baselines (••) also roughy agree with a linear fit to the "selective sampling" points. We therefore hypothesize that **all these methods are mostly addressing label shift**. We verify this hypothesis with the remarkable fit of an additional point (\blacktriangle) of a model trained by resampling according to the <u>test</u> set class distribution, i.e. cheating.

It represents an upper bound that might be achievable in future work with methods for label shift (Azizzadenesheli et al., 2019; Lipton et al., 2018).

We replicated these observations on every test domain of this dataset (Figure 15 in the appendix).



Figure 8: Main results on *civilComments*.

For the above reasons, it makes sense that resampling for uniform classes or combinations greatly improves performance, as shown in prior work (Idrissi et al., 2022).

With selective mixup (), some criterion (same domain/diff. class) performs clearly above all others. But it works even better without mixup! () Among many other variations, none surpasses the uniform-combinations baseline.



treatments. The target metric is the area under the ROC curve (AUROC) which gives equal importance to both classes. We report the worst-domain AUROC, i.e. the lowest value across test time periods.

Vanilla mixup performs a bit better than ERM. Because of the class imbalance, resampling for uniform classes also improves ERM. As expected, this is perfectly equivalent to the selective sampling criterion "diffClass" and they perform therefore equally well. Adding mixup is yet a bit better, which suggests again that **the performance of selective mixup is merely the result of the independent effects of vanilla mixup and resampling**. We further verify this explanation with the novel combination of

simple resampling and vanilla mixup, and observe almost no difference whether the mixing operation is performed or not (last two rows in Figure 9).



To further support the claim that these methods mostly address label shift, we report in Table 1 the proportion of the majority class in the training and test data. We observe that the distribution sampled by the best training methods brings it much closer to that of the test data.

Proportion of majority class	(%)
In the dataset (training)	78.2
In the dataset (validation)	77.8
In the dataset (OOD test)	66.5
Sampled by different training methods	
Resampling (uniform classes)	50.0
Diff. domain + diff. class	50.0
Diff. class	50.1
Same domain + Diff. class	49.9
Resampling (uniform cl.) + concatenated pairs	64.3
Resampling (uniform cl.) + vanilla mixup	64.3

Table 1: The performance of the various methods on *MIMIC-Readmission* is explained by their correction of a class imbalance. The best training methods (boxed numbers) sample the majority class in a proportion much closer to that of the test data.

4.6 EVIDENCE OF A "REGRESSION TOWARD THE MEAN" IN THE DATA

We hypothesized in Section 3 that resampling helps because of a "regression toward the mean" between training and test splits. We now check for this property and find indeed a shift toward uniform class distributions in all datasets studied. For the Wild-Time datasets, we plot in Figure 10 the ratio of the minority class (for binary tasks: *yearbook*, *MIMIC*) and class distribution entropy (for the multiclass task: *arXiv*). Finding this property agrees with the proposed explanation and with the fact that we selected all three datasets because they previously showed improvements with selective mixup in Yao et al. (2022a).

The shift toward uniformity also holds in *waterbirds* and *civilComments*, artificially through the worst-group metric. The training data contains imbalanced groups (class/domain combinations) while the worst-group accuracy gives uniform importance to all groups.



Figure 10: The class distribution shifts toward uniformity in these Wild-Time datasets. This agrees with the explanation that the benefits from resampling rely on a "regression toward the mean".

5 RELATED WORK

Mixup and variants. Mixup was originally introduced in Zhang et al. (2017) and numerous variants followed (Cao et al., 2022). Many propose modality-specific mixing operations: CutMix (Yun et al., 2019a) replaces linear combinations with collages of image patches, Fmix (Harris et al., 2020) combines image regions based on frequency contents, AlignMixup (Venkataramanan et al., 2022)

combines images after spatial alignment. Manifold-mixup (Verma et al., 2019) replaces the mixing in input space with the mixing of learned representations, making it applicable to text embeddings.

Mixup for OOD generalization. Mixup has been integrated into existing techniques for domain adaptation (DomainMix (Xu et al., 2020)), domain generalization (FIXED (Lu et al., 2022b)), and with meta learning (RegMixup (Pinto et al., 2022)). This paper focuses on variants we call "selective mixup" that use non-uniform sampling of the pairs of mixed examples. LISA (Yao et al., 2022b) proposes two heuristics, same-class/different-domain and vice versa, used in proportions tuned by cross-validation on each dataset. Palakkadavath et al. (2022) use same-class pairs and an additional objective to encourage invariance of the representations to the mixing. CIFair (Tian et al., 2023) uses same-class pairs with a contrastive objective to improve algorithmic fairness. SelecMix (Hwang et al., 2022) proposes a selection heuristic to handle biased training data: same class/different biased attribute, or vice versa. DomainMix (Xu et al., 2020) uses different-domain pairs for domain adaptation. DRE (Li et al., 2023) uses same-class/different-domain pairs and regularize their Grad-CAM explanations to improve OOD generalization. SDMix (Lu et al., 2022a) applies mixup on examples from different domains with other improvements to improve cross-domain generalization for activity recognition.

Explaining the benefits of mixup has invoked regularization (Zhang et al., 2020) and augmentation (Kimura, 2021) effects, the introduction of label noise (Liu et al., 2023), and the learning of rare features (Zou et al., 2023). These works focus on the mixing and in-domain generalization, whereas we focus on the selection and OOD generalization.

Training on resampled data. We find that selective mixup is sometimes equivalent to training on resampled or reweighted data. Both are standard tools to handle distribution shifts in a domain adaptation setting (Japkowicz, 2000; Kubat et al., 1997) and are also known as importance-weighted empirical risk minimization (IW-ERM) (Shimodaira, 2000; Gretton et al., 2009). For covariate shifts, IW-ERM assigns each training point x of label y a weight equal to the likelihood ratio $p_{\text{target}}(x)/p_{\text{source}}(x)$, and for label shifts, $p_{\text{target}}(y)/p_{\text{source}}(y)$ (Azizzadenesheli et al., 2019; Lipton et al., 2018), Several works recently showed that reweighting and resampling are competitive with the state of the art in various OOD (Idrissi et al., 2022; Park et al., 2022; Perrett et al., 2023; Sagawa et al., 2019) and label-shift settings (Garg et al., 2023).

6 CONCLUSIONS AND OPEN QUESTIONS

Conclusions. This paper helps understand selective mixup, which is one of the most successful and general methods for distribution shifts. We showed unambiguously that much of the improvements were actually unrelated to the mixing operation and could be obtained with much simpler, well-known resampling methods. On datasets where mixup does bring benefits, we could then obtain even better results by combining the independent effects of the best mixup and resampling variants.

Limitations. We focused on the simplest version selective mixup as described by Yao et al. (2022b). Many papers combine the principle with modifications to the learning objective (Hwang et al., 2022; Li et al., 2023; Lu et al., 2022a; Palakkadavath et al., 2022; Tian et al., 2023; Xu et al., 2020). Resampling likely plays a role in these methods too but this claim requires further investigation. We evaluated "only" five datasets. Since we introduced simple ablations that can single out the effects of resampling, we hope to see future re-evaluations of other datasets.

Because we picked datasets that had previously shown benefits with selective mixup, we cannot fully verify the predicted failure when there is no "regression toward the mean" in the data. Still, we do present one experiment in Appendix C that convincingly verifies this prediction on *yearbook* by swapping the ID and OOD data.

Finally, this work is not about designing new algorithms to surpass the state of the art. Our focus is on improving the scientific understanding of existing mixup strategies and their limitations.

Open questions. Our results leave open the question of the applicability of selective mixup to real situations. The "regression toward the mean" explanation indicates that much of the observed improvements are accidental since they rely on an artefact of some datasets. In real deployments, distribution shifts cannot be foreseen in nature nor magnitude. This is a reminder of the relevance of Goodhart's law to machine learning (Teney et al., 2020) and of the risk of overfitting to popular benchmarks (Liao et al., 2021).

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APPENDICES

A EXPERIMENTAL DETAILS

We follow prior work on each dataset for the **architectures and hyperparameters** of our experiments. For each dataset, all methods compared use hyperparameters initially validated with the ERM baseline. All experiments use early stopping i.e. recording metrics for each run at the epoch of highest ID or worst-group validation performance (for *Wild-Time* and *waterbirds/civilComments* datasets respectively). Each dataset/method is run with 9 different seeds unless otherwise noted. The bar charts report the average over these seeds and error bars represent \pm one standard deviation.

We noticed considerable **variability in the results reported in prior work**, sometimes for datasets/methods supposedly identical (e.g. resampling baselines on *waterbirds*). Therefore we only make comparisons across results obtained within a unique code base after re-running all baselines in the same setting.

We also found some **issues in existing code** that we could not clear up with their authors despite multiple requests. This includes inconsistent preprocessing and duplicated data in the preprocessing of civilComments in Idrissi et al. (2022), "magic constants" in the implementation of selective mixup (LISA) in Yao et al. (2022b), inappropriate architectures for *MIMIC* in Yao et al. (2022a). We fixed these issues in our codebase. Therefore we refrain from claims or direct comparisons with the absolute state of the art.

Dataset-specific notes:

- On *waterbirds*, we use ImageNet-pretrained ResNet-50 models. The results in the main paper use linear classifiers trained on frozen features. We report similar results with fine-tuned ResNet-50 models in Figure 11.
- On *CivilComments*, we use a standard pretrained BERT. To limit the computational expense for our large number of experiments, we use the BERT-tiny version (2 layers, 2 attention heads, embeddings of size 128). The results in the main paper use linear classifiers on frozen features. We report similar results with fine-tuned models in Figure 17 (using only one seed).
- On *Wild-Time Yearbook*, we train the small CNN architecture described in Yao et al. (2022a) from scratch. In the analysis of Figure 5, we measure the distance between the training and test distributions of inputs (vectorized grayscale images). To do so, we measure the distance between every pair across the two sets. For each test example, we keep the minimum distance (i.e. closest training example), then average these distances over the test set.
- On *Wild-Time arXiv*, we use random subset of 10% of the dataset. We verified on a small number of experiments that this produces very similar results to the full dataset at a fraction of the computational expense.
- On *Wild-Time MIMIC-Readmission*, the baseline transformer architecture proposed in Yao et al. (2022b) seems inappropriate. Its ID and OOD performance is surpassed by random guessing or even by constant predictions of the majority training class. The issue probably went unnoticed because the standard accuracy metric is misleading with imbalanced data (70% ID accuracy of that ERM baseline is worse than chance).

To remedy this, we first switch to the AUROC metric. It gives equal weight to the classes and 50% is then unambiguously equivalent to random chance.

Second, we use a much simpler architecture. We train a "bag of embeddings" where each token (diagnosis/treatment code) is assigned a learned embedding, which are summed across sequences then fed to a linear classifier.

All experiments were run on a single laptop with an Nvidia GeForce RTX 3050 Ti GPU.

B ADDITIONAL RESULTS

We show below results from the main paper while including in-domain (ID), out-of-distribution (OOD) average-domain/average-group, and OOD worst-domain/worst-group performance. The OOD metrics are always strongly correlated across methods and training epochs, but ID and OOD performance sometimes require a trade-off, as noted in Teney et al. (2023).



Figure 11: Results on *waterbirds* (top) with linear classifiers on frozen ResNet-50 features and (bottom) with fine-tuned ResNet-50 models (selected methods only).



Figure 12: Results on yearbook.



Figure 13: Results on *yearbook-reversed* (swapping ID and OOD data) to test the predicted failure mode. The "regression toward the mean" does not hold, therefore the methods that improved OOD performance on the original dataset are now detrimental (methods presented in the same order as Figure 12).



Figure 14: Results on *arXiv*. Interestingly, the methods with selective sampling without mixup are much better than selective mixup on in domain (ID) but worse out of domain (OOD). This shows a clear trade-off between these two objectives.



Figure 15: Same analysis as in Figure 7 of the main paper, performed on every test domain. In all cases, we observe a strong correlation between the improvements in accuracy and the reduction in divergence of the class distribution due to resampling effects.



Figure 16: Results on *civilComments* with linear classifiers on frozen embeddings.



Figure 17: Results on *civilComments* with fine-tuned BERT models (single seed, reduced set of methods). These results are qualitatively identical to those with frozen embeddings above.



Figure 18: Results on *MIMIC-Readmission*.

C TESTING THE PREDICTED FAILURE MODE

The explanations proposed in this paper state that the resampling effects with selective mixup are beneficial because a "regression toward the mean" is present in the datasets. As a corollary, it implies that the effect would be detrimental if the opposite property holds (i.e. increased imbalances in the test data).

We test this prediction on the *yearbook* dataset by switching the ID and OOD data. More precisely, whereas the original dataset uses data from years 1930–1970 as training data and ID test data (as shown in Figure 10), we use this data as the OOD test data. Vice versa for data from years 1970–2010. As a result, the training \rightarrow test label shift is now an increased imbalance rather than a regression toward a uniform one.

The results on *yearbook-reversed* confirm the predicted failure (comparing Figures 12–13). The methods that improved OOD performance on the original dataset are now detrimental.

D PROOF OF THEOREM 3.1

Theorem D.1 (Restating Theorem 3.1). Given a dataset $\mathcal{D} = \{(\boldsymbol{x}_i, \boldsymbol{y}_i)\}_i$ and paired data $\widetilde{\mathcal{D}}$ sampled according to the "different class" criterion, i.e. $\widetilde{\mathcal{D}} = \{(\widetilde{\boldsymbol{x}}_i, \widetilde{\boldsymbol{y}}_i) \sim \mathcal{D} \ s.t. \ \widetilde{\boldsymbol{y}}_i \neq \boldsymbol{y}_j\}$, then the distribution of classes in $\mathcal{D} \cup \widetilde{\mathcal{D}}$ is more uniform than in \mathcal{D} .

Formally, the entropy $\mathbb{H}(\mathbf{p}_{Y}(\mathcal{D})) \leq \mathbb{H}(\mathbf{p}_{Y}(\mathcal{D} \cup \widetilde{\mathcal{D}})).$

Proof. Let us define the shorthands $p \stackrel{\text{def}}{=} \mathbf{p}_{\mathbf{Y}}(\mathcal{D})$ and $\widetilde{p} \stackrel{\text{def}}{=} \mathbf{p}_{\mathbf{Y}}(\widetilde{\mathcal{D}})$.

In \hat{D} , the *i*th class gets assigned, in the expectation, on a proportion of points equal to the proportion of all other classes $j \neq i$ in the original data \hat{D} .

Looking at the individual elements of \tilde{p} , we therefore have, $\forall i = 1 \dots C$:

$$\widetilde{p}_i = \sum_{j \neq i}^C p_j / (C - 1) \tag{6}$$

$$\widetilde{p}_i = \left(1 - p_i\right) / \left(C - 1\right) \tag{7}$$

We will show that every element of \tilde{p} is closer to $\frac{1}{C}$ than the corresponding element of p:

$$|p_i - \frac{1}{C}| \ge |\widetilde{p}_i - \frac{1}{C}| \tag{8}$$

$$\left|\frac{C p_{i}-1}{C}\right| \geq \left|\frac{(1-p_{i})C-(C-1)}{C(C-1)}\right|$$
(9)

$$C p_i - 1| \ge \left| \frac{1 - C p_i}{(C - 1)} \right|$$
 (10)

$$|C p_i - 1| \ge |\frac{C p_i - 1}{(C-1)}|$$
 (11)

Therefore \tilde{p} is closer to a uniform distribution than p, and

$$\mathbb{H}(\boldsymbol{p}) \leq \mathbb{H}(\widetilde{\boldsymbol{p}}) \tag{12}$$

Since $\mathbf{p}_{\mathbf{Y}}(\mathcal{D} \cup \widetilde{\mathcal{D}}) = \left(\mathbf{p}_{\mathbf{Y}}(\mathcal{D}) \oplus \mathbf{p}_{\mathbf{Y}}(\widetilde{\mathcal{D}})\right) / 2$, we also have

$$\mathbb{H}(\boldsymbol{p}) \leq \mathbb{H}((\boldsymbol{p} \oplus \widetilde{\boldsymbol{p}})/2)$$
(13)

$$\mathbb{H}(\mathbf{p}_{\mathbf{Y}}(\mathcal{D})) \leq \mathbb{H}(\mathbf{p}_{\mathbf{Y}}(\mathcal{D} \cup \widetilde{\mathcal{D}}))$$
(14)

with an equality iff $\mathbf{p}_{\mathbf{Y}}(\mathcal{D})$ is uniform.