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011 ABSTRACT

013 Sparse Autoencoders (SAEs) are widely used to steer large language models
014 (LLMs), based on the assumption that their interpretable features naturally enable
015 effective model behavior steering. Yet, a fundamental question remains unan-
016 swered: *does higher interpretability indeed imply better steering utility?* To an-
017 swer this question, we train 90 SAEs across three LLMs (Gemma-2-2B, Qwen-
018 2.5-3B, Gemma-2-9B), spanning five architectures and six sparsity levels, and
019 evaluate their interpretability and steering utility based on SAEBENCH (Karvonen
020 et al., 2025) and AXBENCH (Wu et al., 2025) respectively, and perform a rank-
021 agreement analysis via Kendall’s rank coefficients τ_b . Based on the framework,
022 Our analysis reveals only a relatively weak positive association ($\tau_b \approx 0.298$), in-
023 dicating that interpretability is an insufficient proxy for steering performance. We
024 conjecture the interpretability-utility gap may stem from the selection of SAE fea-
025 tures as not all of them are equally effective for steering. To further find features
026 that truly steer the behavior of LLMs, we propose a novel selection criterion: Δ
027 *Token Confidence*, which measures how much amplifying a feature changes the
028 next token distribution. We show that our method improves the steering perfor-
029 mance of three LLMs by **52.52%** compared to the current best output score-based
030 criterion (Arad et al., 2025). Strikingly, after selecting features with high Δ *Token*
031 *Confidence*, the correlation between interpretability and utility vanishes ($\tau_b \approx 0$),
032 and can even become negative. This further highlights the divergence between
033 interpretability and utility for the most effective steering features.

1 INTRODUCTION

036 As Large Language Models (LLMs) become more widely used in real-world applications, ensur-
037 ing the safety of their outputs is increasingly important (Kumar et al., 2024; Ji et al., 2023; Inan
038 et al., 2023). Reliable and controllable behavior is essential for deploying these LLMs in more
039 situations (Chen et al., 2024). Fine-tuning is the standard way to improve controllability, but it re-
040 quires labeled data, significant training time, and compute resources (Hu et al., 2022; Wang et al.,
041 2025c). This has spawned a series of representation-based interventions, i.e., steering, that guide
042 LLM inference by manipulating internal representations, aiming for faster and more lightweight
043 output control (Turner et al., 2023; Turner et al., 2024; Wang et al., 2025b; Stolfo et al., 2025).

044 However, activation-level edits are often coarse: they mix multiple semantics, a phenomenon called
045 polysemy (Bricken et al., 2023). Recently, Sparse Autoencoders (SAEs) have become a valua-
046 ble tool in the interpretability field. They are trained to actively decompose the hidden states of
047 the LLM into sparse and human readable features (Templeton et al., 2024; Mudide et al., 2025).
048 Their interpretable nature has subsequently spurred research into leveraging SAE features for more
049 precise, concept-level control over model behavior (Ferrando et al., 2025; Chalnev et al., 2024).

050 Despite this progress, a critical question remains unanswered: *does higher interpretability truly*
051 *imply better utility?* Since SAEs are trained to balance reconstruction and sparsity to yield human-
052 readable features (Cunningham et al., 2023; Makelov, 2024; O’Brien et al., 2025), their utility for
053 downstream tasks is not a primary objective. Understanding and characterizing this gap is critical
to enabling more interpretable and effective steering with SAEs, because (i) it clarifies when in-

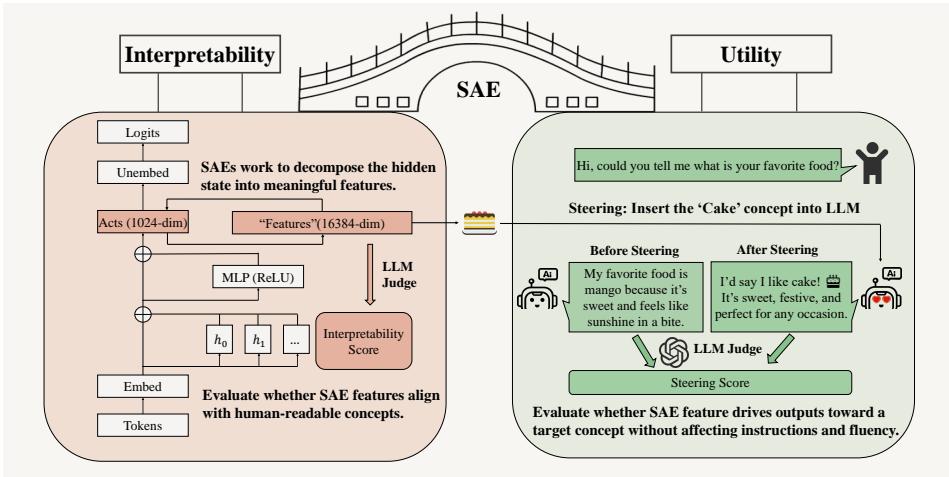


Figure 1: **Overview of our goal: building a bridge for SAE interpretability and utility.** **Interpretability (left):** an SAE attached to the LLM decomposes hidden states into sparse, human-describable features. An LLM judge yields an *interpretability score* for the SAE (Paulo et al., 2025). **Utility (right):** at inference, we activate a target SAE feature (e.g., ‘cake’) to steer generation. An LLM judge yields *steering utility score* (Wu et al., 2025).

interpretability scores can (and cannot) be used as a practical proxy for utility, and (ii) it motivates training objectives that explicitly balance reconstruction, interpretability, and downstream steering performance. To this end, we conduct a systematic study to build a bridge between SAE interpretability and steering utility (see Figure 1).

To perform a comprehensive association analysis, we train 90 SAEs across three LLMs (Gemma-2-2B (Team et al., 2024), Qwen-2.5-3B (Yang et al., 2024), and Gemma-2-9B) spanning diverse architectures and sparsity levels. We compute interpretability using SAEBENCH (Karvonen et al., 2025) and steering utility using AXBENCH (Wu et al., 2025). Then, we leverage a pairwise-controlled framework to evaluate whether interpretability predicts steering performance across the pool of trained SAEs. To quantify this relationship, we follow the idea of prior works (Jiang et al., 2020; Hu et al., 2024) and measure rank agreement between interpretability and utility using Kendall’s rank coefficient τ_b . We control confounders with an axis-conditioned analysis, isolating each design axis (architecture, sparsity, model) by varying one at a time and aggregating per-axis metrics.

Furthermore, as identified in Arad et al. (2025); Wu et al. (2025), not all interpretable features in SAE are equally effective for steering. This motivates our next objective to identify the specific features critical for behavior control and steering utility analysis. Motivated by the recent progress on the entropy mechanism in LLM reasoning (Fu et al., 2025; Wang et al., 2025a), we propose an innovative selection criterion for SAE features: Δ *Token Confidence*, which measures the degree to which amplifying a single feature shifts the model’s next-token distribution. Features that induce the most substantial change in model confidence are identified as high-utility candidates features for steering, as they exert a measurable and targeted influence on model behavior. Finally, we leverage these critical features to conduct a refined analysis of the interpretability-utility gap.

The primary contributions and insights of this paper are summarized as follows:

1. **(§3.4) Interpretability shows a relatively weak positive association with utility.** Across 90 SAEs that are trained across three model sizes, five architectures, and six sparsity levels, we find that a higher *interpretability score* tends to shows a relatively weak positive association with steering performance (the Kendall’s rank coefficient $\tau_b \approx 0.298$). This identifies a notable interpretability-utility gap of the existing SAEs.
2. **(§4.2) Δ Token Confidence effectively selects features with strong steering performance.** To identify the SAE features that are critical for steering, we introduce Δ *Token Confidence*, an innovative metric that identifies steering-critical SAE features by measuring their impact on the model’s next-token distribution. When benchmarked against the best existing output score-based method (Arad et al., 2025), our approach yields a substantial 52.52% average improvement in

108 *steering score*. This result validates the superiority of our method and underscores the critical
 109 role of feature selection in characterizing and enhancing the steering utility of SAEs.

110 3. **(§4.3) The interpretability-utility gap widens among high-utility features.** By reapplying
 111 our association analysis exclusively to SAE features with strong steering utility, we uncover a
 112 counterintuitive finding: the interpretability-utility correlation vanishes or even becomes negative
 113 (Kendall’s rank coefficient $\tau_b \approx 0$). This indicates that for the most effective steering features,
 114 interpretability is at best irrelevant and potentially detrimental, further emphasizing the critical
 115 nature of the interpretability-utility gap.

116 Our results demonstrate a significant divergence between SAE’s interpretability and steering utility,
 117 suggesting that prioritizing interpretability does not enable improved steering performance. This gap
 118 highlights a crucial research direction: mitigating it will likely necessitate advanced post-training
 119 feature selection protocols or fundamentally new, utility-oriented SAE training paradigms.

120

121 2 PRELIMINARY

123 2.1 SPARSE AUTOENCODERS

125 Sparse Autoencoders (SAEs) decompose internal model activations x into sparse, higher-
 126 dimensional features h that can be linearly decoded back to the original space (Cunningham et al.,
 127 2023; Leask et al., 2025). A standard SAE with column-normalized decoder weights (Bricken et al.,
 128 2023; Karvonen et al., 2024) is defined by the following forward map and optimization objective:

129

$$130 \quad \mathcal{L} = \|x - \hat{x}\|_2^2 + \lambda \|h\|_1, \text{ where } h = \text{ReLU}(W_E x + b_E), \hat{x} = W_D h + b_D,$$

132 where W_E, b_E are encoder parameters, W_D, b_D are decoder parameters, \hat{x} is the reconstruction, and
 133 λ controls sparsity. This training balances reconstruction accuracy with sparse representations.

134

135 2.2 INTERPRETABILITY: AUTOMATED INTERPRETABILITY SCORE

136 SAEBENCH (Karvonen et al., 2025) uses an LLM-as-judge (Paulo et al., 2025) to assess each latent:
 137 the judge drafts the description from examples and then predicts, on a held-out set, which sequences
 138 activate it. The *Automated Interpretability Score* is the average precision of the judge’s prediction.

139

$$140 \quad \text{AutoInterp Score} = \frac{1}{M} \sum_{m=1}^M \mathbf{1}[\hat{y}_m = y_m],$$

142

143 where $y_m \in \{0, 1\}$ indicates whether the latent activates in the sequence m and \hat{y}_m is the judge’s
 144 prediction. We use this score as our interpretability metric. For the complete details, see Appendix B.

145

146 2.3 UTILITY: STEERING SCORE

147 SAE steering injects the SAE decoder atom v_f (the f -th column of the column-normalized decoder
 148 $W_{\text{dec}}[f]$) into the residual stream at a target layer to push the hidden state x along a chosen feature
 149 direction (Durmus et al., 2024). Given a feature index f , a steering factor α , and a per-sample scale
 150 m_f (e.g., the feature’s maximum activation), the intervention is

151

$$152 \quad x^{\text{steer}} = x + (\alpha m_f) \cdot v_f. \quad (1)$$

153

154 Through the above formula (1), we can use SAE features for steering to achieve the output of controlling
 155 LLM. AxBENCH (Wu et al., 2025) measures causal control by steering internal representations
 156 during generation and asking an LLM judge to rate three aspects, each on a discrete scale $\{0, 1, 2\}$:
 157 *Concept (C)*, *Instruction (I)*, and *Fluency (F)*. The overall *Steering Score* is the harmonic mean:

158

$$159 \quad \text{Steering Score} = \text{HM}(C, I, F) = \frac{3}{\frac{1}{C} + \frac{1}{I} + \frac{1}{F}} \in [0, 2].$$

160

161

162 Following AxBENCH, for each concept we sample instructions (e.g., 10 from Alpaca-Eval (Dubois
 163 et al., 2023)), generate continuations under different steering factors, pick the best factor on one
 164 split, and evaluate the held-out split with the judge to obtain the final utility score averaged across
 165 prompts (Gu et al., 2025). The complete scoring procedure is detailed in Appendix C.

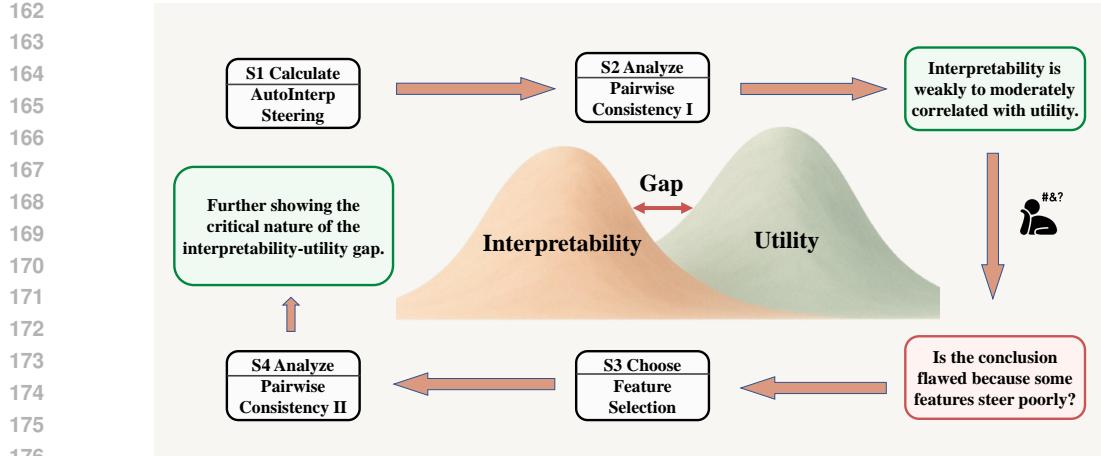


Figure 2: **Overview of our pairwise-controlled workflow linking SAE interpretability with steering utility.** **(S1)** Compute *interpretability score* and *steering score* for each SAE. **(S2)** Pairwise analysis across SAEs and get an insight (the top-right green box), revealing an interpretability-utility gap. The red box (lower right) is our further inference based on the above green box and previous studies (Wu et al., 2025). **(S3)** Use Δ *Token Confidence* to select higher-utility features. **(S4)** Compute *steering score* after selection per SAE, then do the pairwise analysis between *steering score* and *interpretability*. The green box in the middle left is our final conclusion.

3 CAN SAE INTERPRETABILITY INDICATE STEERING PERFORMANCE?

3.1 EXPERIMENTAL SETUP

Dataset. For each trained SAE, we score 1,000 latents with *LLM-as-judge* (Paulo et al., 2025) and randomly sample 100 to form that SAE’s CONCEPT100: 100 human-readable concept descriptions per evaluation set, each pairing a latent (layer, feature_id) with a short natural-language description. (see details in Appendix F). For steering, we sample 10 Alpaca-Eval instructions, allow up to 128 generated tokens, and test 6 steering factors; the 10 instructions are split 5/5 for factor selection vs. held-out evaluation.

Model. We evaluate three open LLMs: Gemma-2-2B (Team et al., 2024), Qwen-2.5-3B (Yang et al., 2024), and Gemma-2-9B (Team et al., 2024). SAEs are trained on residual-stream activations at a fixed mid-layer for each model: Layer 12 for Gemma-2-2B, Layer 17 for Qwen-2.5-3B, and Layer 20 for Gemma-2-9B—and steering is applied to the corresponding layer.

SAE with different architectures We train 90 SAEs covering a range of architectures and sparsity. All SAEs use a latent dictionary width of 16k. We instantiate five variants: BatchTopK (Bussmann et al., 2024), Gated (Rajamanoharan et al., 2024a), JumpReLU (Rajamanoharan et al., 2024b), ReLU (Team, 2024), TopK (Gao et al., 2024) and sweep six target sparsity levels with approximate per-token activations $L_0 \approx 50, 80, 160, 320, 520, 820$. Further details are provided in Appendix A.

3.2 PAIRWISE RANK CONSISTENCY BETWEEN INTERPRETABILITY AND UTILITY

We test whether higher interpretability of SAE is predictive of higher steering performance across a set of trained SAEs attached to a fixed LM. For each SAE θ in a pool Θ , we record a pair $(\mu(\theta), g(\theta)) \in \mathbb{R}^2$, where μ is the SAE-level *Interpretability Score* and g is an aggregated *Steering Score* over a standardized evaluation suite.

Given two SAEs $\theta_i, \theta_j \in \Theta$, define the concordance indicator

$$v_{ij} = \text{sign}(\mu(\theta_i) - \mu(\theta_j)) \cdot \text{sign}(g(\theta_i) - g(\theta_j)) \in \{-1, 0, +1\}. \quad (2)$$

216 Kendall’s tie-corrected rank coefficient τ_b (Kendall, 1938) summarizes agreement over unordered
 217 pairs and reduces to average concordance when there are no ties:
 218

$$\tau_b = \frac{\sum_{i < j} v_{ij}}{\sqrt{\left(\sum_{i < j} \mathbf{1}[\mu(\theta_i) \neq \mu(\theta_j)]\right) \left(\sum_{i < j} \mathbf{1}[g(\theta_i) \neq g(\theta_j)]\right)}} \in [-1, 1], \quad (3)$$

224 In this study, we instantiate μ with the *Interpretability Score* and g with the *Steering Score*, then
 225 compute τ for three model–layer settings (Gemma-2-2B, Qwen-2.5-3B, Gemma-2-9B). Each setting
 226 includes 30 SAEs spanning architectures and sparsity to ensure sufficient pair coverage.
 227

228 3.3 GRANULATED KENDALL’S COEFFICIENT TO CONTROL CONFOUNDERS

230 Global rank agreement can be confounded by hyperparameters that jointly influence interpretability
 231 and utility. To obtain an axis-controlled assessment, we factor the SAE design space into orthogonal
 232 axes and evaluate rank consistency while varying one axis at a time and holding the others fixed.

233 We define three conditioning axes: (A) Architecture — fix architecture (and layer), vary sparsity;
 234 (B) Sparsity — compare architectures at matched sparsity ranks; (C) Model — fix the base model,
 235 compare all SAEs within it. For axis i , partition Θ into groups \mathcal{G}_i that are matched on all axes except
 236 i . Within each group $G \in \mathcal{G}_i$, compute Kendall’s coefficient in $\{(\mu(\theta), g(\theta)) : \theta \in G\}$, and average
 237 between groups to obtain the statistic at the axis level:

$$\psi_i = \frac{1}{|\mathcal{G}_i|} \sum_{G \in \mathcal{G}_i} \tau(\{(\mu(\theta), g(\theta)) : \theta \in G\}). \quad (4)$$

241 Aggregate the axis-level outcomes by

$$\Psi = \frac{1}{n} \sum_{i=1}^n \psi_i, \quad (5)$$

242 where n is the number of axes. Each ψ_i captures rank consistency conditioned on axis i (varying only
 243 that axis while matching the others), and Ψ aggregates these into a single axis-controlled measure.
 244 This construction mitigates cross-axis trends—e.g., architecture, sparsity, or model-driven shifts that
 245 can obscure local relationships between interpretability and utility.

246 We report the per-axis statistics ψ_i together with the aggregate Ψ for the same model settings as in
 247 section 3.2, providing both axis-specific and aggregated assessments.
 248

249 3.4 PAIRWISE ANALYSIS RESULTS

250 In this section, we assess whether higher SAE interpretability predicts stronger steering by computing
 251 Kendall’s τ_b between the *Interpretability Score* $\mu(\theta)$ and the *Steering Score* $g_{\text{base}}(\theta)$ (before any
 252 feature selection) over a pooled set of SAEs attached to a fixed LLM.
 253

254 To control confounders and localize effects, we apply the axis-conditioned procedure defined in
 255 section 3.3. For each axis, we form matched groups, compute within-group τ_b , average to obtain a
 256 per-axis summary, and aggregate these summaries into an overall axis-controlled coefficient.
 257

258 Table 1 shows that **across SAEs, higher interpretability tends to be modestly associated with**
 259 **better steering on average, pointing to a consistent but limited impact.** The pooled Kendall’s
 260 $\tau_b \approx 0.30$ is positive, and the axis-controlled aggregate remains positive ($\Psi \approx 0.25$), indicating that
 261 more interpretable features generally translate into better steering utility across designs and models.
 262

263 **The strength of the link between interpretability and utility depends on SAE architecture,**
 264 **sparsity, and the base model.** By architecture, the association is positive on average ($\Psi_A \approx 0.26$),
 265 with ReLU-like variants reinforcing the trend and Gated weakening it. By sparsity, alignment is
 266 strongest when the SAE is more sparse and weakens—sometimes reversing—as the number of active
 267 features increases. By model, the underlying LM shapes the effect, with the signal clearest in Qwen-
 268 2.5-3B and weaker in Gemma-2-2B, while the model-wise summary remains positive ($\Psi_C \approx 0.33$).
 269

270
271 **Table 1: Pairwise Analysis Between Interpretability Score $\mu(\theta)$ and Steering Score $g_{\text{base}}(\theta)$.** Here $g_{\text{base}}(\theta)$
272 is steering score before any feature selection. We report Kendall’s τ_b overall and by axis-controlled measures
273 Ψ_A (Architecture), Ψ_B (Matched Sparsity), and Ψ_C (Model). n = number of SAEs; Pairs = number of pairwise
274 comparisons; p = permutation p -value for testing $H_0 : \tau_b = 0$ within each group (by shuffling steering scores);
275 95% CI = 95% bootstrap percentile confidence interval for τ_b .

Axis	SAEs	n	Pairs	τ_b	p	95% CI
Overall	All SAEs	90	4005	0.2979	0.0002	[0.1591, 0.4191]
$\Psi_A = \mathbf{0.2575}$ (SE ≈ 0.1163 , 95% boot CI [0.0222, 0.3961])						
	BatchTopK	18	153	0.3203	0.0712	[0.0000, 0.6429]
	Gated	18	153	-0.2026	0.2577	[-0.5493, 0.1918]
Ψ_A : Architecture	JumpReLU	18	153	0.4248	0.0160	[0.0563, 0.7183]
	ReLU	18	153	0.3595	0.0392	[0.0274, 0.6528]
	TopK	18	153	0.3856	0.0272	[0.0282, 0.6966]
$\Psi_B = \mathbf{0.1651}$ (SE ≈ 0.1112 , 95% boot CI [-0.0286, 0.3587])						
	$L_0 \approx 50$	15	105	0.5429	0.0034	[0.1578, 0.8737]
	$L_0 \approx 80$	15	105	0.3524	0.0740	[0.0000, 0.6304]
Ψ_B : Sparsity	$L_0 \approx 160$	15	105	0.1810	0.3821	[-0.2501, 0.5464]
	$L_0 \approx 320$	15	105	0.1810	0.3673	[-0.1579, 0.5208]
	$L_0 \approx 520$	15	105	-0.2190	0.2837	[-0.5914, 0.1648]
	$L_0 \approx 820$	15	105	-0.0476	0.8484	[-0.4409, 0.3750]
$\Psi_C = \mathbf{0.3272}$ (SE ≈ 0.0698 , 95% boot CI [0.2184, 0.4575])						
Ψ_C : Model	Gemma-2-2B	30	435	0.2184	0.0980	[-0.0644, 0.4764]
	Qwen-2.5-3B	30	435	0.4575	0.0008	[0.2086, 0.6580]
	Gemma-2-9B	30	435	0.3057	0.0166	[0.0521, 0.5277]
$\Psi = (\Psi_A + \Psi_B + \Psi_C)/3 = \mathbf{0.2499}$						

296 **Key Observation 1:** Interpretability shows a relatively weak positive correlation with steering
297 performance, highlighting a notable gap between interpretability and utility across SAEs.

4 FROM INTERPRETABILITY TO UTILITY: WHICH SAE FEATURES ACTUALLY STEER?

303 In Sec. 3.4, We find that SAE interpretability is a relatively weak prior for steering utility. Prior
304 work (Arad et al., 2025) shows many features lack steerability and we speculate that this factor may
305 render the previous conclusion inaccurate. Therefore, we introduce a metric to identify steering-
306 effective features. Metrics derived from a model’s internal token distributions can assess reasoning
307 quality (Kang et al., 2025). In particular, *token entropy* offers a unified view: high entropy highlights
308 critical decision points (Fu et al., 2025; Wang et al., 2025a). We apply this idea to SAE steering.

4.1 FEATURE SELECTION VIA Δ TOKEN CONFIDENCE

311 We start from the model’s next-token distribution. Given logits $z \in \mathbb{R}^V$ and $p = \text{softmax}(z)$ over a
312 vocabulary of size V , the *token entropy* is

$$H(p) = - \sum_{j=1}^V p_j \log p_j, \quad (6)$$

317 Entropy summarizes dispersion over the vocabulary: smaller values reflect a sharper, more concentrated
318 prediction, while larger values indicate greater uncertainty at a given position.

319 To focus on the head of the distribution that matters most for sampling, we use *token confidence*.
320 Let $\mathcal{I}_k(p) \subseteq \{1, \dots, V\}$ denote the indices of the k largest probabilities in p . The top- k *token
confidence* is the negative average log-probability over these entries:

$$C_k(p) = -\frac{1}{k} \sum_{j \in \mathcal{I}_k(p)} \log p_j. \quad (7)$$

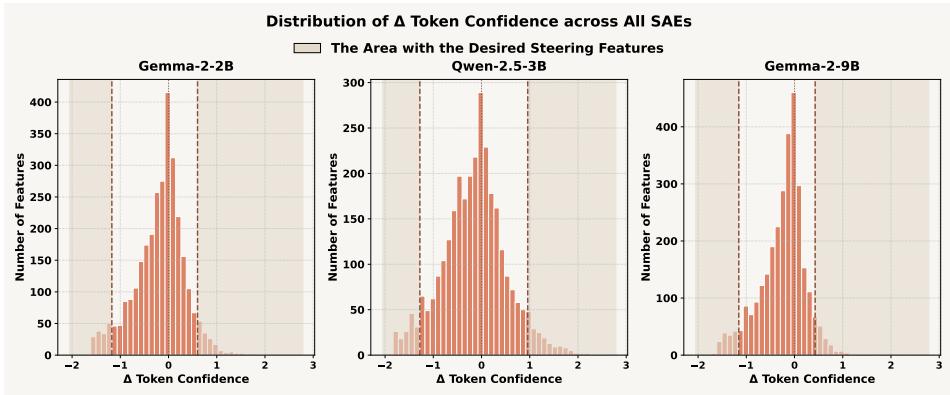


Figure 3: **Distribution of per-feature Δ Token Confidence across all SAEs.** Panels show histograms for Gemma-2-2B, Qwen-2.5-3B, and Gemma-2-9B; the x -axis is ΔC_k (negative values indicate increased confidence, positive values decreased confidence) and the y -axis is the number of SAE features. The shaded area marks the high-magnitude tails from which candidate steering features are selected, while the central mass near 0 indicates features with little distributional impact.

Lower C_k implies higher confidence, while higher C_k implies a flatter top- k distribution. Unlike entropy, C_k directly captures the sharpness of the outcomes that drive next-token behavior.

We turn confidence into a feature-level selector via a single-feature SAE intervention. Consider an SAE feature f at layer ℓ . We amplify only the coefficient of f by a factor $\alpha > 0$ in the SAE reconstruction, leaving the base model and all other features unchanged. Denote the baseline next-token distribution by p^{base} and the intervened distribution by $p_{f,\ell,\alpha}^{\text{int}}$. Δ Token Confidence is

$$\Delta C_k(f; \ell, \alpha) = C_k(p_{f,\ell,\alpha}^{\text{int}}) - C_k(p^{\text{base}}). \quad (8)$$

Negative values $\Delta C_k < 0$ mean that amplifying f sharpens the top- k distribution, while positive values indicate greater dispersion. We compute this using one baseline and one intervened forward pass via an SAE hook. Implementation details and hyperparameters are provided in Appendix D.

We select features with extreme signed changes in *token confidence* under single-feature SAE interventions, and pick the better direction (see Figure 3). For each feature, we compute ΔC_k , rank by $|\Delta C_k|$, form tiers, evaluate subsets for steering, and keep the best per SAE.

4.2 STEERING PERFORMANCE RESULTS AFTER FEATURE SELECTION

Arad et al. (2025) has shown that SAE steering works well if features are chosen by their causal impact on model outputs, introducing the *output score* as a metric to identify output-aligned features. For a given SAE feature, a logit-lens procedure is used to select a representative token set M . Let $P_{\text{base}}(M)$ and $P_{\text{int}}(M)$ denote the aggregate support for M in the base and intervened (feature-amplified) runs, respectively, where

$$P(M) = \left(1 - \frac{\min_{i \in M} \text{rank}(i)}{|V|}\right) \max_{i \in M} p(i),$$

with $|V|$ the vocabulary size, $\text{rank}(i)$ the rank of token i , and $p(i)$ its probability. The single-feature steering score is then

$$S_{\text{out}} = P_{\text{int}}(M) - P_{\text{base}}(M),$$

which is large when amplifying the feature increases the rank and probability of representative token set. Following this insight, we evaluate our Δ token confidence selection on three base LLMs (Gemma-2-2B, Qwen-2.5-3B, Gemma-2-9B) using the CONCEPT100 (see details in 3.1). We set $k = 1$ in our Δ token confidence selection. In ablations over $k \in \{1, 3, 5, 10\}$ on Gemma and Qwen, all choices of k yield large gains over the SAE baseline, and $k = 1$ is in fact the best-performing setting on both models, so we adopt it as our default (see Appendix L). The experiments on steering performance improvement of each SAE can be referred to Appendix E.

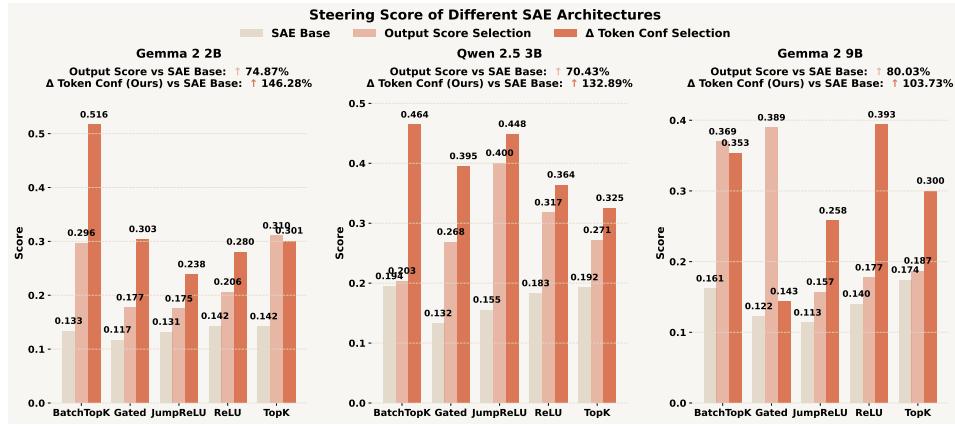


Figure 4: **Comparison of different SAE steering methods with five SAE architecture across three LLMs.** Panels correspond to Gemma-2-2B, Qwen-2.5-3B, and Gemma-2-9B. The horizontal axis groups SAE architectures (BatchTopK, Gated, JumpReLU, ReLU, TopK), and the vertical axis reports the *steering score*. Bars show three conditions: SAE Base (no feature selection), Output Score Selection, and Δ Token Confidence Selection (ours). Panel annotations summarize the average lift of each selection method relative to the SAE-based steering.

Table 2 shows that our selection yields consistent gains across all models, outperforming the vanilla SAE baseline by large margins, and also improving over an output-score-based selector. These gains indicate that ranking and filtering by the magnitude of distributional change captured by ΔC_k reliably isolates features with the strongest steering utility.

Furthermore, we conducted a comparative analysis of SAEs of different architectures on three models. For fair comparison, the two feature selection methods use the same subset size. Figure 4 compares *steering scores* across SAE architectures and selection methods. In all three models, selecting features by Δ Token Confidence consistently outperforms both the no-selection SAE baseline and the output-score selector across architectures.

On average, our method improves steering performance by **52.52%** over the strongest competing baseline. The BatchTopK architecture is the one that has the most stable and significant improvement in steering capabilities on models of different sizes among the five SAE architectures.

Key Observation 2: Δ Token Confidence reliably selects high-utility SAE features across models. Among SAE architectures, BatchTopK achieves the most stable and sizable *steering score*.

4.3 PAIRWISE ANALYSIS AFTER HIGH-CONFIDENCE FEATURE SELECTION

Building on the high Δ Token Confidence feature selection introduced above, we now ask whether SAE interpretability can serve as a prior for *post-selection steering performance*. For each SAE θ , we write $g_{\text{base}}(\theta)$ for its *Steering Score* before feature selection and $g_{\text{high}}(\theta)$ for the best score achieved after selecting features with high Δ Token Confidence. Both quantities are measured in the same steering metric; $g_{\text{high}}(\theta)$ is simply the steering score recomputed on a restricted set of high Δ Token Confidence features for the same SAE.

432

433 **Table 3: Pairwise Analysis Between Interpretability Score $\mu(\theta)$ and Steering Score $g_{\text{high}}(\theta)$.** Here $g_{\text{high}}(\theta)$
 434 is steering score after selecting features with high Δ Token Confidence. We report Kendall’s τ_b overall and
 435 under axis-controlled summaries Ψ_A (Architecture), Ψ_B (Matched Sparsity), and Ψ_C (Model). n = number of
 436 SAEs; Pairs = number of pairwise comparisons; p = permutation p -value for testing $H_0 : \tau_b = 0$ within each
 437 group (by shuffling steering scores); 95% CI = 95% bootstrap percentile confidence interval for τ_b .

Axis	SAEs	n	Pairs	τ_b	p	95% CI
Overall	All SAEs	90	4005	0.0823	0.2645	[-0.0560, 0.2243]
$\Psi_A = \mathbf{0.0618}$ (SE ≈ 0.1136 , 95% boot CI $[-0.1149, 0.2810]$)						
	BatchTopK	18	153	-0.1895	0.2933	[-0.5694, 0.1918]
	Gated	18	153	0.0338	0.8764	[-0.3656, 0.3853]
Ψ_A : Architecture	JumpReLU	18	153	0.0950	0.6209	[-0.3373, 0.4697]
	ReLU	18	153	-0.1007	0.5803	[-0.4833, 0.2857]
	TopK	18	153	0.4702	0.0078	[0.0794, 0.8043]
	$\Psi_B = \mathbf{0.1002}$ (SE ≈ 0.0700 , 95% boot CI $[-0.0288, 0.2188]$)					
Ψ_B : Sparsity	$L_0 \approx 50$	15	105	0.2488	0.2212	[-0.1828, 0.6300]
	$L_0 \approx 80$	15	105	0.2899	0.1602	[-0.0673, 0.5876]
	$L_0 \approx 160$	15	105	0.1183	0.5747	[-0.2357, 0.4151]
	$L_0 \approx 320$	15	105	0.1171	0.5803	[-0.3531, 0.5209]
	$L_0 \approx 520$	15	105	-0.1827	0.3747	[-0.6490, 0.3524]
	$L_0 \approx 820$	15	105	0.0097	1.0000	[-0.3724, 0.3572]
Ψ_C : Model	$\Psi_C = \mathbf{0.0538}$ (SE ≈ 0.0700 , 95% boot CI $[-0.0540, 0.1850]$)					
	Gemma-2-2B	30	435	-0.0540	0.6963	[-0.3127, 0.2125]
	Qwen-2.5-3B	30	435	0.1850	0.1668	[-0.0872, 0.4413]
	Gemma-2-9B	30	435	0.0303	0.8284	[-0.2552, 0.3233]
$\Psi = (\Psi_A + \Psi_B + \Psi_C)/3 = \mathbf{0.0719}$						

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457

458 We quantify the relationship between interpretability and post-selection steering by computing
 459 Kendall’s τ_b between the *Interpretability Score* $\mu(\theta)$ of SAEs and the *Steering Score* $g_{\text{high}}(\theta)$. As
 460 in Section 3.4, we report both pooled coefficients and axis-conditioned summaries that control for
 461 design and model factors (architecture, sparsity and model). In addition, Section 3.4 and Table 1
 462 report the corresponding coefficients for $\mu(\theta)$ vs. $g_{\text{base}}(\theta)$, allowing a direct comparison between
 463 “before” and “after” feature selection using the same steering metric.

464 Overall, Table 3 shows that interpretability is not a reliable prior for steering performance after
 465 selection. The pooled association between $\mu(\theta)$ and $g_{\text{high}}(\theta)$ is small and statistically indistinguishable
 466 from zero ($\tau_b \approx 0.08$), and the axis-controlled aggregate is likewise close to zero ($\Psi \approx 0.07$).
 467 Estimates across architectures, sparsity levels, and models cluster around zero and are mostly non-
 468 significant. These results indicate that once we focus on features that are most useful for steering (as
 469 selected by high Δ Token Confidence), higher interpretability can’t predict better steering scores.

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Key Observation 3: Surprisingly, the interpretability–utility gap widens after feature selection:
 higher interpretability scores can’t provide a prior for which SAEs achieve better steering.

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5 RELATED WORK

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5.1 REPRESENTATION-BASED STEERING

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Activation-based steering arose as a lightweight alternative to fine-tuning, enabling on-the-fly control of LLM behavior without retraining (Giulianelli et al., 2018; Vig et al., 2020; Geiger et al., 2021; 2025). The core idea is to inject carefully chosen directions into hidden states, typically in the residual stream, scaling interventions by a gain and selecting layers for maximal effect (Zou et al., 2025; Rimsky et al., 2024; van der Weij et al., 2024). It has been applied to safety and moderation, persona and sentiment control, and instruction adherence, promising low-latency deployment-time adjustment but facing polysemantic entanglement and brittleness that motivate standardized evaluation (Chen et al., 2025; Liu et al., 2024). However, this approach injects polysemantic activations at intervention time, yielding coarse-grained effects for output control (Bricken et al., 2023). Our work

486 is related to activation-level interventions, but differs by grounding directions in sparse, interpretable
 487 SAE features and applying utility-oriented feature selection to mitigate these failure modes.
 488

489 **5.2 SAE-BASED STEERING**
 490

491 Sparse Autoencoders (SAEs) decompose activations into sparse, human-readable features to mit-
 492 igate polysemy and expose concept-level structure (Bricken et al., 2023; Templeton et al.,
 493 2024; Gao et al., 2024). For steering, practitioners use decoder atoms as directions and add scaled
 494 injections at chosen layers, with architecture and sparsity choices trading reconstruction for feature
 495 granularity (Zhao et al., 2025; Wang et al., 2025d; Ferrando et al., 2025). SAE-based steering en-
 496 ables targeted safety control, style modulation, and instruction emphasis, yet the utility of individual
 497 features varies widely (Chalnev et al., 2024; Mayne et al., 2024). While the connection between
 498 SAE interpretability and steering utility remains unclear, and our goal is to build a principled bridge
 499 between them. To this end, we conduct a large-scale experiments across multiple model sizes and
 500 SAE architectures, demonstrating the critical nature of the interpretability-utility gap.
 501

502 **6 CONCLUSION AND DISCUSSION**
 503

504 In summary, SAE interpretability shows relatively weak positive association with steering utility
 505 across 90 SAEs ($\tau_b \approx 0.298$), revealing a clear interpretability-utility gap. Selecting features with
 506 Δ *Token Confidence* yields substantial gains (average +52.52% over the strongest existing baseline).
 507 Surprisingly, when analyzing *steering score* after selection, the correlation with interpretability col-
 508 lapses toward zero and can even turn negative for the highest-utility features, further underscoring
 509 this gap. This gap points to a key direction: develop task-general utility indicators that reliably
 510 predict steerability across models, or design training objectives that directly optimize controllability
 511 under sparsity so features are utility-calibrated without heavy post-hoc selection. Our work provides
 512 valuable insight for the further development of SAEs as interpretable tools.
 513

514 **REPRODUCIBILITY STATEMENT**

515 We aim to facilitate full reproduction of our results. All model code, training and evaluation scripts,
 516 configuration files, and experiment logs are released at an *anonymous* repository as part of the sup-
 517 plementary materials: <https://anonymous.4open.science/r/SAE4Steer>. Training architectures, hy-
 518 perparameters, sparsity schedules, and optimization details are specified in the main text and Ap-
 519 pendix A.2 (see also the per-family settings in Appendix A). The datasets used are openly licensed:
 520 all SAEs are trained on The Common Pile v0.1 (Kandpal et al., 2025) as described in Appendix F;
 521 our evaluation concepts (CONCEPT100) and their automatic generation pipeline are documented
 522 in Appendix B and Appendix F. The complete procedures for automated interpretability scoring
 523 (SAEBENCH) and steering utility (AXBENCH), including sampling, judging protocols, and scoring
 524 functions, are detailed in Appendix B and Appendix C, with the Δ *Token Confidence* selector defined
 525 in Appendix D and the post-selection results summarized in Appendix E. Hardware, runtime, and
 526 memory footprints for both SAEBENCH and AXBENCH are reported in Appendices B.3 and C.2.
 527 Together, these materials, along with seed-controlled configuration files and exact command-line
 528 invocations provided in the anonymous repository, are intended to enable independent researchers
 529 to replicate and extend our findings.
 530

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790 LLM USAGE

791

792 In preparing this paper, large language models (LLMs) were used as an assistive tool for minor
 793 language polishing and stylistic improvements. All technical contributions, results, and conclusions
 794 are solely the work of the authors.

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797 A SAE ARCHITECTURES AND TRAINING DETAILS

798

799 We train **90 SAEs** (30 per base model) across five architectures and six target sparsity levels. Unless
 800 stated otherwise, the dictionary width is **16K** codes ($F=16,384$), SAEs are attached to the residual
 801 stream at the layer described in the main text, and decoder columns are ℓ_2 -normalized. All models
 802 are trained on The Common Pile v0.1 (Kandpal et al., 2025).
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806 A.1 ARCHITECTURES AND PARAMETERIZATION

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808 We list the five SAE families with their named parameters (as implemented) and the correspond-
 809 ing shapes. The last column records architecture-specific thresholding/gating fields when present.

Shapes assume residual dimension $d=2304$ and dictionary width $F=16,384$.

810 811 812 813 814 815 816 817 818 819 820 821 822 823 824 825 826 827 828 829 830 831 832 833 834 835 836 837 838 839 840 841 842 843 844 845 846 847 848 849 850 851 852 853 854 855 856 857 858 859 860 861 862 863	Architectures	W_{enc}	b_{enc}	W_{dec}	b_{dec}	Threshold / Extras
ReLU	encoder.weight: shape (16,384, 2,304)	encoder.bias: shape (16,384)	decoder.weight: shape (2,304, 16,384)	decoder.bias: shape (2,304)	—	
Gated	encoder.weight: shape (16,384, 2,304)	gate_bias: shape (16,384)	decoder.weight: shape (2,304, 16,384)	decoder_bias: shape (2,304)	r_mag: shape (16,384); mag_bias: shape (16,384)	
TopK	encoder.weight: shape (16,384, 2,304)	encoder.bias: shape (16,384)	decoder.weight: shape (2,304, 16,384)	b_dec: shape (2,304)	k	
BatchTopK	encoder.weight: shape (16,384, 2,304)	encoder.bias: shape (16,384)	decoder.weight: shape (2,304, 16,384)	b_dec: shape (2,304)	k	
JumpReLU	W_{enc} : shape (2,304, 16,384)	b_{enc} : shape (16,384)	W_{dec} : shape (16,384, 2,304)	b_{dec} : shape (2,304)	threshold: shape (16,384)	

A.2 TRAINING, SPARSITY, AND COMPUTE SETUP

Optimization and schedule. Adam with learning rate 3×10^{-4} ; LR warmup 1000 steps; sparsity warmup 5000 steps; LR decay starting at 80% of total steps. Precision: bfloat16. LM batch size = 4, context length = 2048, SAE batch size = 2048. Each run trains on $\sim 5 \times 10^8$ tokens.

Sparsity controls. We sweep six target activity levels

$$L_0 \approx \{50, 80, 160, 320, 520, 820\}.$$

For TopK/BatchTopK we set k equal to the chosen L_0 (aux- k coefficient 1/32; moving-threshold momentum 0.999; threshold tracking begins at step 1000). JumpReLU uses the same set via target_10. For L_1 -penalized families, we search the following penalty grids:

845 846 847 848 849 850 851 852 853 854 855 856 857 858 859 860 861 862 863	Family	L_1 penalty values (used to span sparsity levels)
Standard / Standard-New		0.012, 0.015, 0.020, 0.030, 0.040, 0.060
Gated SAE		0.012, 0.018, 0.024, 0.040, 0.060, 0.080

Training details. All training uses two NVIDIA RTX A800 GPUs. The table below reports the aggregated artifacts and training time (hours) for 30 SAEs *per model* (total 90), together with the runtime configuration. Times and sizes are approximate.

855 856 857 858 859 860 861 862 863	Model	#SAEs	Disk (GB)	Traing Time (H)	LM Batch	Context	SAE Batch	Peak Mem (GB)
Gemma-2-2B	30	8.7	17	4	2048	2048	20	
Gemma-2-9B	30	13.2	60	4	2048	2048	70	
Qwen-2.5-3B	30	7.7	37	4	2048	2048	30	

864 **B SAEBENCH DETAILS, RESULTS AND OUR COSTS**
865866 **B.1 AUTOMATED INTERPRETABILITY SCORE PROCESS**
867868 SAEBENCH (Karvonen et al., 2025) follow an LLM-as-judge pipeline to assign an *automated in-*
869 *terpretability score* to each SAE latent. First, we collect layer activations by running the base LM
870 with caching and encoding the residual stream through the SAE to obtain $h \in \mathbb{R}^{N \times L \times F}$. We define
871 a token window of length 21 (buffer = 10) around any center (i, t) and, unless stated otherwise,
872 mask BOS/PAD/EOS positions. For a latent ℓ , we sample three window types: (i) **Top** ($n = 12$
873 non-overlapping peaks of $h[:, :, \ell]$), (ii) **Importance-Weighted** ($n = 7$, sampled proportional to
874 activation after removing values at least as large as the smallest Top peak), and (iii) **Random** ($n = 10$,
875 uniform over valid centers). Let v_{\max} be the maximum activation seen in any Top window position
876 and set a global threshold $\tau_{\text{act}} = 0.01 v_{\max}$.
877878 We split the sampled windows into a **generation set** (10 Top + 5 IW) and a **scoring set** (2 Top + 2
879 IW + 10 Random, shuffled). In generation, tokens with activation $> \tau_{\text{act}}$ are bracketed to highlight
880 evidence; the judge LLM receives these 15 windows and returns a short English description of when
881 the latent fires. In scoring, the judge sees the description and the 14 held-out windows *without*
882 highlights and outputs a comma-separated list of indices it predicts as activations (or `None`).
883884 Ground truth for a window \mathcal{W} is $\mathbb{1}[\max_{u \in \mathcal{W}} h[u, \ell] > \tau_{\text{act}}]$. The per-latent score is the **accuracy**
885 over the $M = 14$ scoring windows, i.e.,
886

887
$$\text{Score}(\ell) = \frac{1}{M} \sum_{m=1}^M \mathbf{1}[\hat{y}_m = y_m],$$

888 where $\hat{y}_m \in \{0, 1\}$ is the judge prediction and y_m is the label defined above. For each SAE θ , we
889 evaluate 1,000 latents and report the mean over a random CONCEPT100 subset:
890

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$$\mu(\theta) = \frac{1}{100} \sum_{\ell \in \text{CONCEPT100}} \text{Score}(\ell).$$

892

893 **B.2 PERFORMANCE OF SAEs ON THREE MODELS ON SAEBENCH**
894895 Across the three backbones, the six SAEBENCH metrics (for information about these indicators, see
896 SAEBENCH (Karvonen et al., 2025)) jointly reveal how sparsity mechanisms balance interpretabil-
897 ity, faithfulness, and causal structure. Automated Interpretability is strongest when encoders enforce
898 compact latent usage (e.g., TopK/BatchTopK and ReLU at lower L_0), and it gradually softens as ca-
899 pacity expands. The Absorption metric (considered via its complement in the plots) indicates that de-
900 signs concentrating signal into a small set of latents are less prone to feature stealing, whereas higher
901 effective capacity encourages redundancy and competition across latents. Meanwhile, Core/Loss-
902 Recovered remains uniformly high, showing that even sparse codes closely preserve original model
903 behavior; increasing L_0 pushes faithfulness toward a ceiling without overturning the core trade-offs
904 visible in the other metrics.
905906 **Gemma-2-2B.** As shown in Fig. 5, Gemma-2-2B exhibits a balanced profile: interpretability stays
907 robust for TopK/BatchTopK and ReLU at modest sparsity; absorption is contained when the code
908 remains compact; and Core is near-saturated across the range. Improvements in SCR@20 are steady
909 but measured, suggesting targeted debiasing with small k . Sparse Probing indicates that relatively
910 few latents already carry much of the predictive signal, while RAVEL strengthens with moderate
911 capacity, reflecting cleaner separation of attributes without undermining compactness.
912913 **Qwen-2.5-3B.** For Qwen-2.5-3B (Fig. 6), interpretability at low-to-moderate L_0 is competi-
914 tive—especially for TopK and JumpReLU—yet the model is more sensitive to absorption as ca-
915 pacity grows, implying greater latent competition and signal spread. Core remains excellent, so
916 reconstructions are faithful; however, SCR gains can flatten at high L_0 where residual spurious cues
917 reappear. Sparse Probing is solid but a touch behind the strongest Gemma configurations, consist-
918 ent with its flatter RAVEL patterns: causal structure is present but less crisply disentangled when
919 attributes begin to diffuse across latents.
920

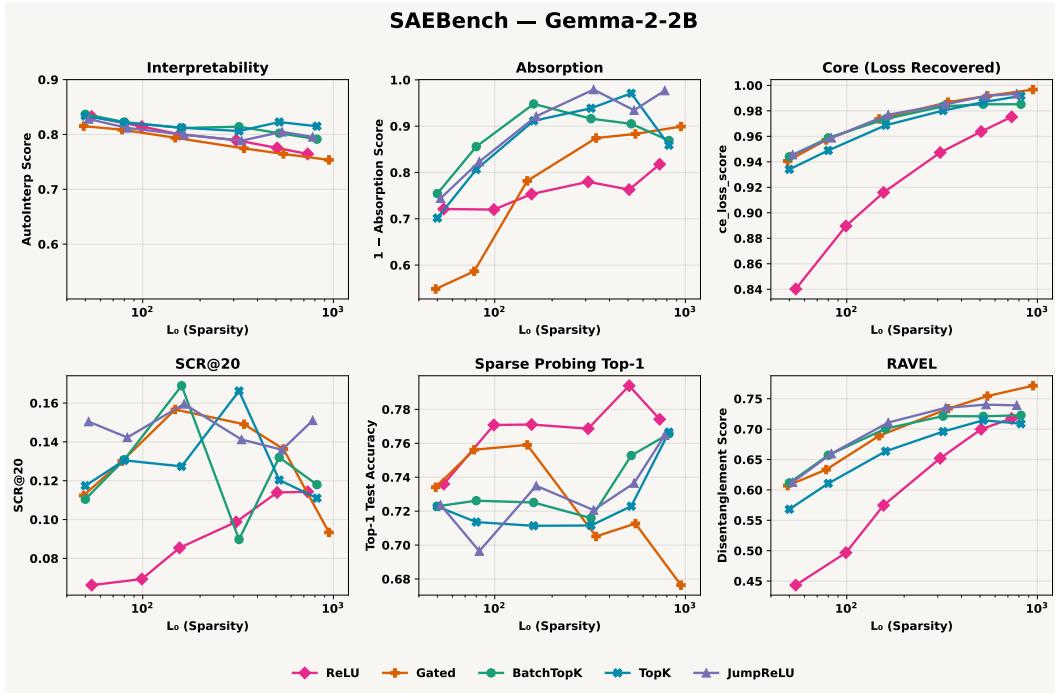


Figure 5: SAEbench results for **Gemma-2-2B**: interpretability remains strong at lower L_0 , absorption stays low for compact codes, Core is near ceiling, and structure (SCR/RAVEL) improves with moderate capacity.

Gemma-2-9B. Gemma-2-9B (Fig. 7) pushes the upper envelope on structure: interpretability remains solid for compact encoders; absorption is low at moderate L_0 that avoids unnecessary latent proliferation; and Core is near its ceiling. SCR@20 is the most decisive among the three, pointing to cleaner isolation of spurious factors with small, targeted ablations. Sparse Probing is strong and, together with higher RAVEL, indicates that only a handful of latents capture both predictive signal and causally specific attributes with minimal collateral interference.

B.3 SAEBENCH RUNTIME COST

The computational requirements for running SAEbench evaluations were measured on two NVIDIA RTX A800 GPUs using **16K**-width SAEs trained on the Gemma-2-2B (Team et al., 2024), Qwen-2.5-3B (Yang et al., 2024) and Gemma-2-9B. Table 4 summarizes the *per-SAE* runtime for each evaluation type. Several evaluations include a one-time setup phase (e.g., precomputing activations or training probes) that can be reused across multiple SAEs; after this setup, each evaluation has its own runtime per SAE. We therefore report amortized per-SAE minutes.

Table 4: **Approximate SAEbench runtime per SAE (minutes).** Values are per-SAE and represent amortized minutes after any one-time setup; each minute figure is an approximation and may vary with hardware and I/O.

Model	Core	Interpretability	Absorption	Sparse Probing	Ravel	SCR
Gemma-2-2B 4	8	12	2	18	10	
Qwen-2.5-3B 7	9	15	8	17	16	
Gemma-2-9B 11	12	17	30	40	28	

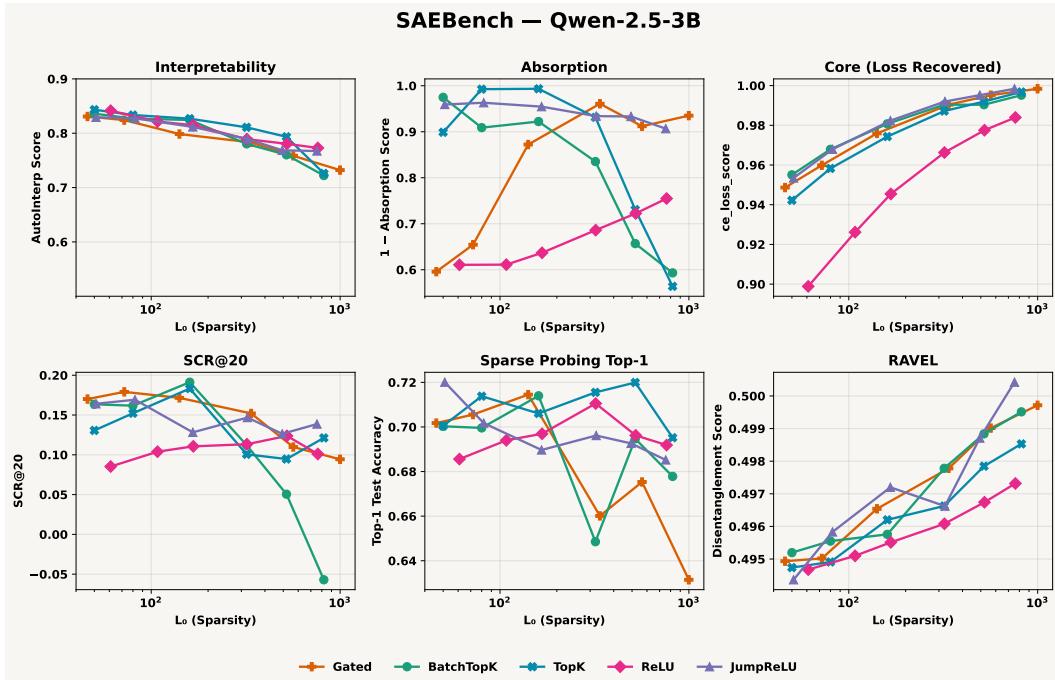


Figure 6: SAEbench results for **Qwen-2.5-3B**: strong interpretability at lower L_0 , rising absorption with capacity, consistently high Core, and more fragile SCR/RAVEL at the highest capacities.

C AXBENCH DETAILS AND OUR COSTS

C.1 STEERING SCORE PROCESS

For each SAE feature (concept) we evaluate steering utility with an AxBENCH-style judge (Wu et al., 2025). Given the concept description from SAEbench and the intervention rule in Eq. equation 1, we sample 10 instructions (Alpaca-Eval style) and split them 5/5 into a development set for factor selection and a held-out set for evaluation. We generate continuations (up to 128 tokens) for a grid of steering factors $\alpha \in \mathcal{A}$ by injecting the feature direction at the target layer. For each generated continuation, an LLM judge assigns three subscores on the discrete scale $\{0, 1, 2\}$ (floats allowed but clamped to $[0, 2]$): *Concept* (C ; does the output express the concept?), *Instruction* (I ; is it on-topic for the instruction?), and *Fluency* (F ; is it well-formed?). We aggregate with the harmonic mean

$$\text{HM}(C, I, F) = \begin{cases} 0, & \text{if } C \leq 0 \text{ or } I \leq 0 \text{ or } F \leq 0, \\ \frac{3}{\frac{1}{C} + \frac{1}{I} + \frac{1}{F}}, & \text{otherwise,} \end{cases} \in [0, 2].$$

On the development split we pick $\alpha^* = \arg \max_{\alpha \in \mathcal{A}} \frac{1}{5} \sum \text{HM}$, then evaluate the held-out split with α^* to obtain the *per-concept Steering Score* (mean HM over the 5 held-out prompts). SAE-level utility $g(\theta)$ is the average of per-concept scores over the evaluation set for that SAE.

CONCEPT TEMPLATE

[System]

Please act as an impartial judge and evaluate whether the specified concept is presented in the sentence fragment provided below. Focus solely on whether the concept is clearly incorporated, without regard for grammar or logical coherence.

Begin your evaluation by providing a short explanation of whether the concept is clearly incorporated. Point out references to the concept in the fragment, noting any minimal or awkward incorporations. After providing your explanation, rate the concept's relevance on a scale from 0 to 2, where 0 indicates the concept is not present at all, 1 indicates the concept is somewhat present but minimally or awkwardly incorporated, and 2 indicates the concept is more fully and effectively incorporated, with stronger and

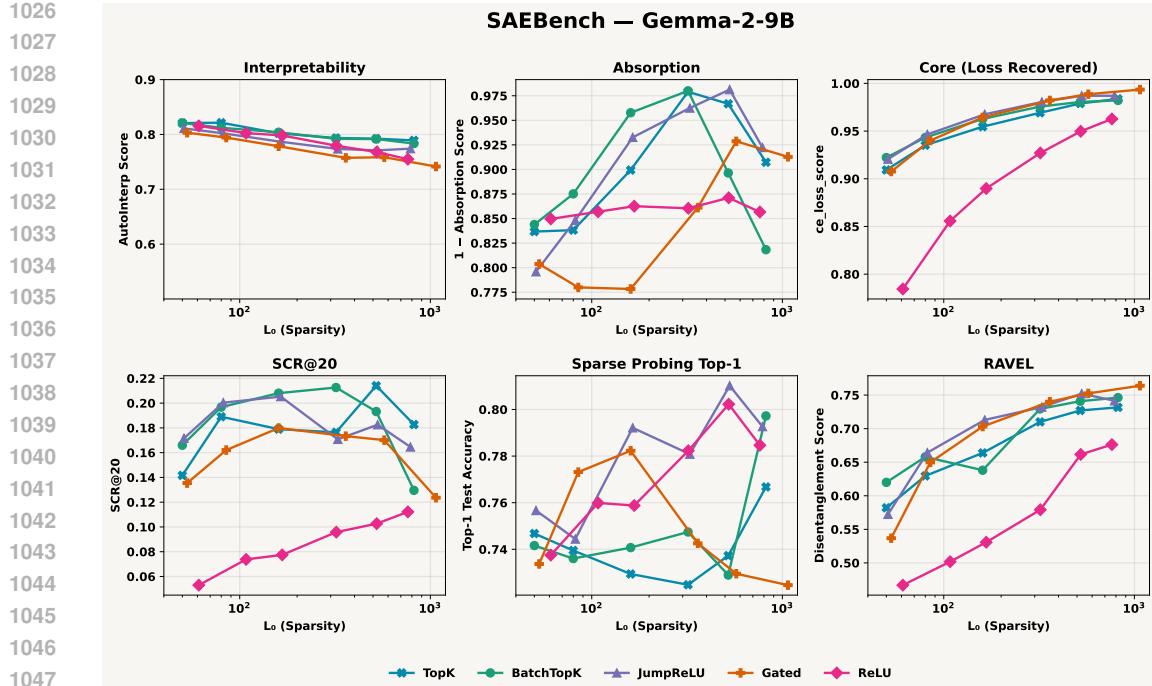


Figure 7: SAEbench results for **Gemma-2-9B**: robust interpretability with compact codes, low absorption at moderate L_0 , near-ceiling Core, and the clearest gains in SCR/RAVEL among the three backbones.

more natural integration. Provide your rating using this exact format: “Rating: [[score]]”.

[Concept Start] {concept} [Concept End]

[Sentence Fragment Start] {sentence} [Sentence Fragment End]

INSTRUCT_TEMPLATE

[System]

Please act as an impartial judge and evaluate whether the sentence fragment provided below is related to the instruction. Focus solely on the degree of relatedness in terms of topic, regardless of grammar, coherence, or informativeness.

Begin your evaluation by providing a brief explanation of whether the sentence is related to the instruction, and point out references related to the instruction. After providing your explanation, rate the instruction relevance on a scale from 0 to 2, where 0 indicates the sentence is unrelated to the instruction, 1 indicates it is somewhat related but only minimally or indirectly relevant in terms of topic, and 2 indicates it is more clearly and directly related to the instruction. Provide your rating using this exact format: “Rating: [[score]]”.

[Instruction Start] {instruction} [Instruction End]

[Sentence Fragment Start] {sentence} [Sentence Fragment End]

FLUENCY_TEMPLATE

[System]

Please act as an impartial judge and evaluate the fluency of the sentence fragment provided below. Focus solely on fluency, disregarding its completeness, relevance, coherence with any broader context, or informativeness.

Begin your evaluation by briefly describing the fluency of the sentence, noting any unnatural phrasing, awkward transitions, grammatical errors, or repetitive structures that may hinder readability. After providing your explanation, rate the sentence’s fluency on a scale from 0 to 2, where 0 indicates the sentence is not fluent and highly unnatural (e.g., incomprehensible or repetitive), 1 indicates it is somewhat fluent but contains noticeable errors or awkward phrasing, and 2 indicates the sentence is fluent and almost perfect. Provide your rating using this exact format: “Rating: [[score]]”.

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[Sentence Fragment Start] {sentence} [Sentence Fragment End]

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C.2 AXBENCH STEERING EVALUATION COST

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All steering-score evaluations were run on **two NVIDIA RTX A800 GPUs**. The LLM-as-judge backend was `gpt-4o-mini`. Evaluating one SAE on CONCEPT100 costs approximately **\$5** in judge API fees; with **90** SAEs total (≈ 30 per model), the per-model API cost is about **\$150**. Table 5 lists approximate per-SAE runtime and peak VRAM for each model.

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Table 5: AxBench steering evaluation cost per SAE. Runtimes are per-SAE (hours) and approximate; VRAM is peak memory (GB). Judge fees assume `gpt-4o-mini`: $\sim \$5$ per SAE on CONCEPT100; -Model Cost assumes ~ 30 SAEs/model ($\approx \$150$).

Model	Runtime / SAE (h)	Peak VRAM (GB)	Per-Model Cost (USD)
Gemma-2-2B	15	10	150
Qwen-2.5-3B	16	12	150
Gemma-2-9B	23	36	150

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D IMPLEMENTATION OF Δ TOKEN CONFIDENCE

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For a fixed, neutral prefix s (we use “*From my experience,*”, following the previous work(Arad et al., 2025)) we compare the next-token distribution of the base model with that of an intervened model in which a single SAE feature is amplified at layer L by a factor α . The intervention is applied via the same SAE hook point used during training (on the residual stream of block L). We then compute the change in a confidence surrogate built from the top- k probabilities.

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Token confidence. (Fu et al., 2025) For a distribution p over the vocabulary, let $p_{(1)} \geq \dots \geq p_{(k)}$ be the top- k probabilities.

$$C_k(p) = -\frac{1}{k} \sum_{i=1}^k \log p_{(i)}.$$

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Delta token confidence. With p_{base} from a standard forward pass and p_{int} from a pass with the SAE feature intervention,

$$\Delta C_k(f; \alpha, L) = C_k(p_{\text{int}}) - C_k(p_{\text{base}}).$$

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Each feature f is evaluated with two single-step forwards on the same prefix s : (i) a baseline pass; (ii) an intervened pass where we scale the code for f by α before decoding it into the residual at layer L while keeping all other codes at zero. Hooks are removed immediately after the intervened pass to prevent accumulation across evaluations. In this work, we choose $\alpha = 10$ and $k = 1$.

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Feature selection from ΔC_k . For each SAE we rank its features by ΔC_k in two directions: *UP* (largest positive ΔC_k) and *DOWN* (most negative ΔC_k). We form selection sets using either (i) top- K by magnitude with $K \in \{1, 2, 3, 4, 5\}$ per direction, or (ii) upper/lower-tail quantiles (e.g., $q \in \{0.99, 0.95, 0.90, 0.80\}$ mirrored for the lower tail). These sets are then carried into AXBENCH (Wu et al., 2025) to measure utility lift.

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E STEERING RESULTS OF SAE ARCHITECTURES AFTER FEATURE SELECTION

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We quantify steering with the AXBENCH judge after selecting features using Δ Token Confidence (Appendix D). Unless otherwise noted, lifts are reported as the percentage change of a given SAE’s steering score relative to its own baseline (no selection). Results are organized at three levels: aggregate across SAEs per base model, per-SAE rankings, and distribution by architecture.

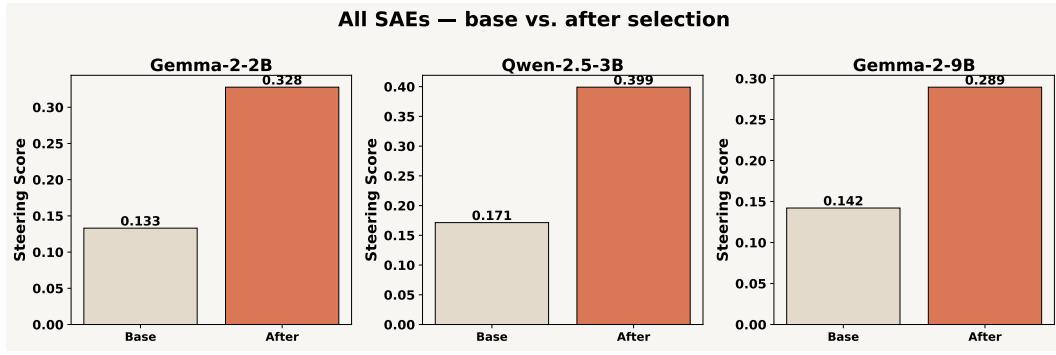


Figure 8: **Overall steering score before and after feature selection.** For each base model, the panel shows two bars: the average baseline *steering score* across its SAEs and the average after applying Δ Token Confidence-based selection. Bars are annotated with the corresponding values; axes share the same scale across panels.

The aggregate view in Figure 8 summarizes how selection affects the mean *steering score* across all SAEs of a base model. Using Δ Token Confidence for feature selection markedly improves the *steering score* across all three models in the figure. For Gemma-2-2B, the score rises from 0.133 to 0.328, which is a 146.6% relative improvement. Qwen-2.5-3B increases from 0.171 to 0.399, a 133.3% improvement, and achieves the highest post-selection score overall. Gemma-2-9B moves from 0.142 to 0.289, a 103.5% improvement.

Conclusions: (i) feature selection via Δ Token Confidence consistently boosts steering for all models; (ii) relative gains are largest for the smallest model (Gemma-2-2B) and smallest for the largest model (Gemma-2-9B), suggesting diminishing relative returns with scale; and (iii) in absolute terms, Qwen-2.5-3B reaches the strongest final *steering score* after selection.

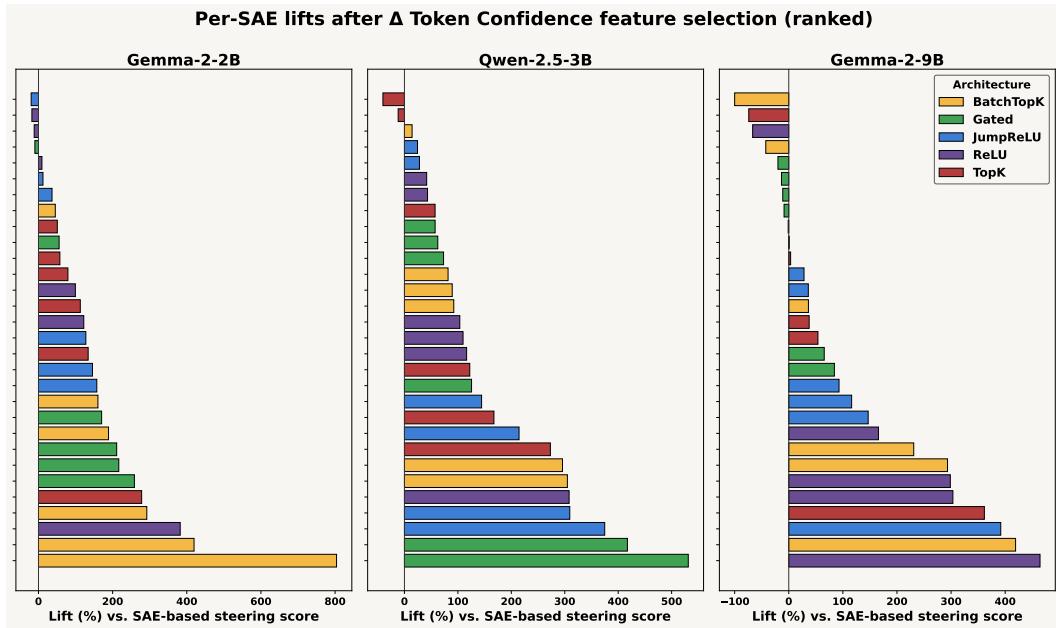
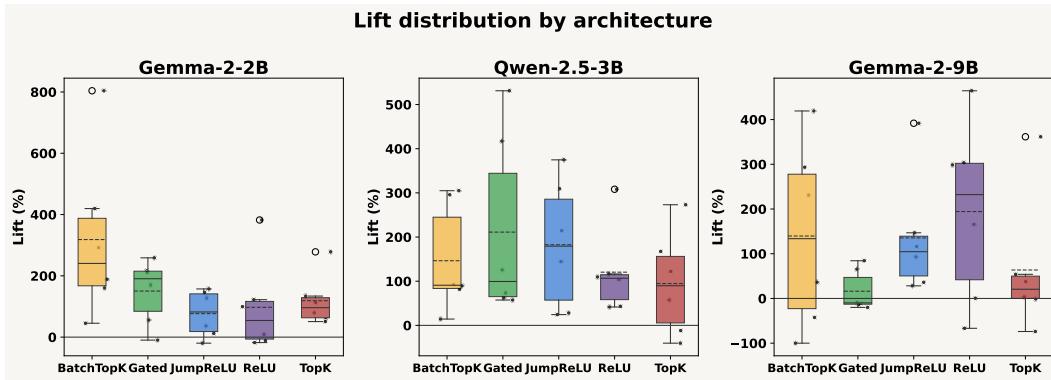


Figure 9: **Per-SAE percentage lift after Δ Token Confidence selection.** Each panel corresponds to a base model. Horizontal bars report the percent lift of the SAE-level *steering score* relative to its own baseline, sorted from largest to smallest within the panel. Bar colors indicate the SAE training architecture (legend shared across panels).

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 1189 Figure 9 ranks SAEs within each model by their relative lift. Architecturally, no single SAE training
 1190 approach dominates; however, the top-ranked lifts are frequently occupied by *BatchTopK* and
 1191 *Gated* variants, with *ReLU/JumpReLU* also contributing strongly and *TopK* showing more mixed
 1192 outcomes. Overall, Δ *Token Confidence* yields consistent per-SAE gains, with variance decreasing
 1193 and stability increasing as model size grows, while architectural diversity remains valuable for
 1194 capturing the largest lifts.



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 1208 Figure 10: **Lift distributions by SAE architecture.** For each base model, box-and-whisker plots
 1209 (with individual points overlaid) summarize the distribution of percentage lifts grouped by training
 1210 architecture. Dashed horizontal lines denote the mean within each group, and whiskers follow the
 1211 conventional interquartile rule.

1212 Figure 10 groups lifts by architecture to visualize differences in central tendency and dispersion
 1213 under the same selection and evaluation protocol. Read together with the per-SAE ranking, this
 1214 distributional view helps disentangle architecture effects from model-specific variation and
 1215 indicates which families tend to produce more stable or more variable outcomes after feature selection.
 1216 *BatchTopK* and *Gated* generally occupy the highest central tendency with wide—but mostly positive—
 1217 spread, especially on Gemma-2-2B and Qwen-2.5-3B. *BatchTopK* achieves the most stable
 1218 and sizable *steering score*. Variance is largest for the smallest model (Gemma-2-2B), indicating
 1219 architecture-sensitive wins at small scale.

F DATASET

1223 **Training corpus for SAEs.** We train all SAEs on **The Common Pile v0.1** (Kandpal et al., 2025),
 1224 an openly licensed ~ 8 TB text collection built for LLM pretraining from ~ 30 sources spanning
 1225 research papers, code, books, encyclopedias, educational materials, and speech transcripts. The
 1226 corpus was curated as a principled alternative to unlicensed web text and has been validated by
 1227 training competitive 7B models on 1–2T tokens. We use it as the sole pretraining dataset for all SAE
 1228 runs. More training details provided in Appendix A.2.

1229 **CONCEPT100 for steering utility.** To evaluate steering, we construct CONCEPT100: a
 1230 compact benchmark of 100 *human-readable concept descriptions* per evaluation set, pro-
 1231 duced automatically by our interpretability pipeline (Appendix B). Each entry is a pair
 1232 (`layer_feature_id, description`) that summarizes a latent’s activation pattern in plain lan-
 1233 guage (e.g., mathematical symbols, scientific terms, pronouns, or domain phrases). These descrip-
 1234 tions are supplied to the AXBENCH judge when computing *steering score*. The examples below
 1235 illustrate the style and domain coverage.

Gemma-2-9B, BatchTopK, $L_0 \approx 80$. Ten examples:

1. 20_14429: concepts related to optical communication systems and their performance characteristics
2. 20_5795: specific technical terms and chemical compounds often related to scientific contexts

1242 3. 20_7908: terms related to gravitational lensing and its effects in cosmology
 1243 4. 20_3042: pronouns and verbs indicating relationships or contributions in various contexts
 1244 5. 20_11897: scientific measurements and units related to energy, concentration, or bi-
 1245 ological data
 1246 6. 20_12944: terms related to cell types and apoptosis mechanisms in scientific contexts
 1247 7. 20_8796: references to specific authors and statistical concepts in mathematical contexts
 1248 8. 20_6430: the phrase “action” in mathematical and theoretical contexts
 1249 9. 20_2220: chemical elements and compounds, particularly including metals and
 1250 metal-related terms
 1251 10. 20_585: various forms of the word “energy” and related concepts in scientific contexts

1252 **Qwen-2.5-3B, Gated, $L_0 \approx 72$.** Ten examples:

1253 1. 17_15113: terms related to mathematical concepts and various scientific names or
 1254 terms
 1255 2. 17_11476: the phrase “as a function of” in contexts of measurement and analysis
 1256 3. 17_162: mathematical symbols and concepts related to coordinates, magnitudes, and
 1257 parameters in equations
 1258 4. 17_2552: dataset identifiers and technical terms common in research and academic
 1259 documents
 1260 5. 17_16377: mathematical notation and technical terms commonly found in formal
 1261 documents
 1262 6. 17_3195: demographic, clinical, and biological characteristics in study populations
 1263 and related comparisons
 1264 7. 17_9186: specific technical terms and concepts related to networking and program-
 1265 ming
 1266 8. 17_11487: mathematical notation and variables related to functions and equations
 1267 9. 17_14657: mathematical notations and structures involving angle brackets and prop-
 1268 erties of functions
 1269 10. 17_1256: terms related to errors and error correction in coding theory and quantum
 1270 operations

1271 **Gemma-2-2B, JumpReLU, $L_0 \approx 81$.** Ten examples:

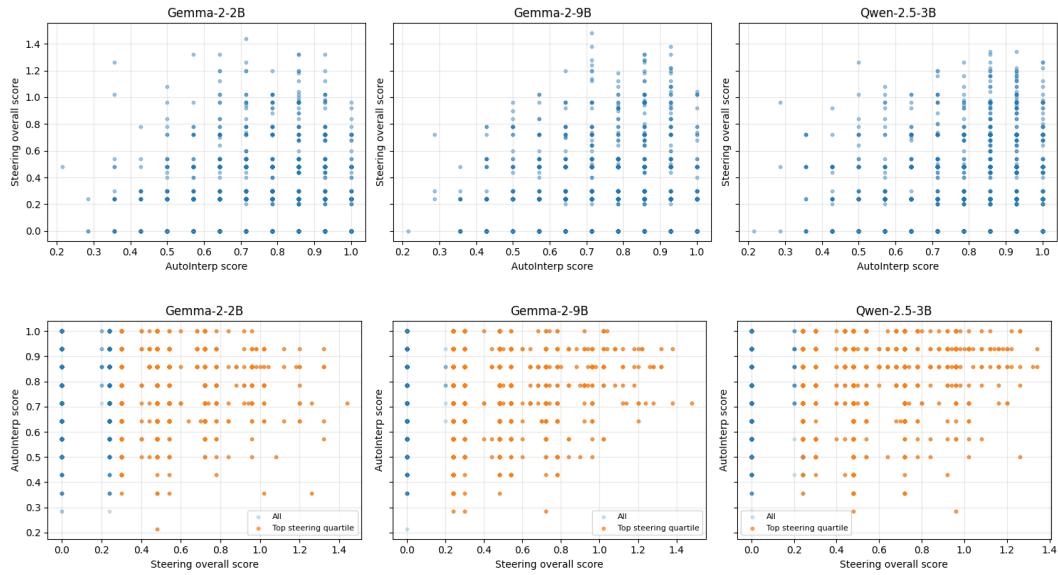
1272 1. 20_11531: terms related to sports, programming, or specific keywords from various
 1273 contexts
 1274 2. 20_10460: terms related to fractional differential equations and numerical methods
 1275 for solving them
 1276 3. 20_4882: terms related to asymptotic theory, robustness, and statistical estimation
 1277 methods
 1278 4. 20_4425: first-person plural pronouns and expressions of intention or conjecture
 1279 5. 20_372: technical terms related to measurement and structure in scientific contexts
 1280 6. 20_9999: the word “from” and contexts implying deviation or distance from some-
 1281 thing
 1282 7. 20_9703: the word “based” in various contexts of theoretical foundations and
 1283 methodologies
 1284 8. 20_15509: technical or numerical concepts in a variety of contexts

1296
 1297 9. 20_8218: phrases indicating conditions or assumptions that must be met in theoretical
 1298 contexts
 1299 10. 20_4614: time intervals and durations mentioned in the context of studies or obser-
 1300 vations

1301 We currently maintain 90 SAEs (30 per base model). Beyond the CONCEPT100 sets evaluated in
 1302 this paper, we have constructed the CONCEPT1000 and CONCEPT16K suites that scale the number
 1303 of human-readable concepts up to 16K. We will extend training and evaluation to these larger suites
 1304 in forthcoming releases to further substantiate the reliability and generality of this work.
 1305

1306 G ANALYSIS OF STEERING STRENGTH ACROSS INTERPRETABILITY LEVELS

1308 To directly address the concern that highly interpretable SAE features might have negligible steering
 1309 effects, we conduct an additional analysis combining interpretability scores with steering scores at
 1310 the feature level.



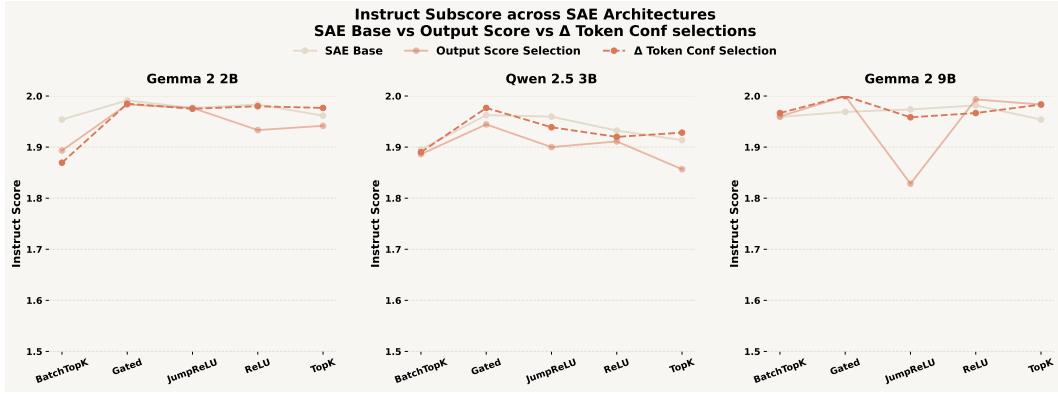
1331 **Figure 11: Feature-level relationship between interpretability and steering strength across**
 1332 **models.** Panels are shown separately for Gemma-2-2B, Gemma-2-9B, and Qwen-2.5-3B. The first
 1333 set of panels plots interpretability score (AutoInterp) against steering overall score for every latent.
 1334 The second set of panels flips the axes and highlights the top steering quartile in orange, with all
 1335 latents shown in blue.

1336 We do not find that interpretable SAE features have negligible steering effects. As the first set of
 1337 panels in Figure 11 shows, high-interpretability features span the full range of steering strengths,
 1338 including many strongly steering latents; high-interpretability features are not dominated by weak
 1339 steering. Conversely, the second set of panels in Figure 11 shows that strongly steering features
 1340 (top steering quartile) also cover almost the full range of interpretability scores, and many of them
 1341 are only moderately or weakly interpretable according to interpretability score. In the SAEs under
 1342 Qwen-2.5-3B, we found that features with extremely high steering scores also tend to have higher
 1343 interpretability scores. This explains why, in Appendix J, only the SAE model under Qwen-2.5-3B
 1344 shows improved steering performance when using high interpretability scores. However, below 1.0,
 1345 the performance becomes highly dispersed again.

1346 Taken together, these two views indicate that (i) high interpretability does not guarantee strong
 1347 steering, and (ii) strong steering does not require high interpretability. In other words, feature-level
 1348 interpretability scores provide limited information about how strongly a latent direction pushes the
 1349 logits, reinforcing our claim that interpretability and steering strength are nearly orthogonal at the
 feature level.

1350 1351 H INSTRUCTION-FOLLOWING UNDER FEATURE-SELECTION STRATEGIES

1352 To address whether our feature-selection metric surfaces features that make model outputs incoherent, we compare the AxBench *Instruction* subscore when steering with three strategies: using all
 1353 SAE features (“SAE Base”), using features selected by the Output Score baseline, and using features
 1354 selected by our Δ Token Confidence metric. For each model and SAE architecture, we aggregate
 1355 the instruction subscores of the steered generations and plot them in Figure 12.
 1356



1357
 1358 **Figure 12: Instruction subscore across SAE architectures when steering with different feature-
 1359 selection strategies.** For each model, we compare the AxBench *Instruction* score under three setups:
 1360 using all SAE features (SAE Base), using Output Score selection, and using Δ Token Confidence
 1361 selection.
 1362

1363 Across all models, both selection methods keep the instruction subscore close to the base level (typ-
 1364 ically in the 1.5–2.0 range on a 0–2 scale), indicating that neither Output Score nor Δ Token Con-
 1365 fidence systematically degrades instruction following. This behavior is expected because coherence
 1366 is primarily controlled by the *steering factor* α in the intervention
 1367

$$x^{\text{steer}} = x + (\alpha m_f) \cdot v_f,$$

1368 which scales how strongly a feature f is injected into the hidden state. AxBench already tunes α
 1369 per feature to balance concept shift against instruction following and fluency. The role of Δ Token
 1370 Confidence is not to push the model into incoherent regimes, but to select features for which even
 1371 small interventions induce a strong, directional change in the token-level logits.
 1372

1373 I EXPLORING THE STEERING UTILITY OF SAEs WITH HIGH 1374 INTERPRETABILITY

1375 To directly test whether selecting features by interpretability also selects for high steering utility, we
 1376 run a controlled comparison at the feature level. For each SAE in our three model–layer settings, we
 1377 compute three steering configurations on AxBench: (i) **SAE Base**, which steers with all evaluated
 1378 features; (ii) **Interpretability selection**, which steers using only the top- $p\%$ latents by AutoInterp
 1379 score (we use $p = 10\%$ in Figure 13); (iii) **Δ Token Confidence selection**
 1380

1381 Across all three models, selecting features by interpretability alone yields at best modest and often
 1382 negative changes in steering strength relative to using all SAE features (about 35% for Gemma-2-2B,
 1383 +10% for Qwen-2.5-3B, and 19% for Gemma-2-9B), whereas Token Confidence consistently deliv-
 1384 ers large gains of roughly 1.1–1.5 \times over the same SAE base. This pattern indicates that feature-level
 1385 interpretability is a weak signal for steering utility, while Token Confidence reliably concentrates
 1386 probability mass on features that produce substantially stronger steering effects.
 1387

1388 J SIGNIFICANCE ANALYSIS FOR DIFFERENT SAE STEERING METHODS

1389 To address the question of whether the improvements in Figure 4 are statistically significant, we
 1390 reran the analysis at the SAE level and, for each model and architecture, computed the mean steering
 1391

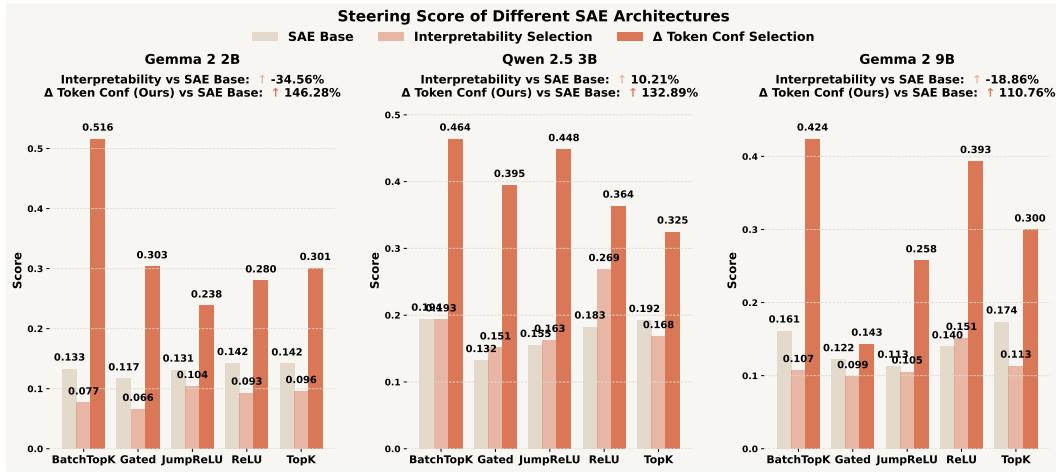


Figure 13: **Steering scores under different feature-selection strategies across SAE architectures and three models.** Bars show the AxBench overall steering score for (left) steering with all features (SAE Base), (middle) steering with the top-10% most interpretable features (AutoInterp), and (right) steering with features selected by Δ Token Confidence. The banners report the macro-average percentage lift over the SAE Base.

score and its 95% confidence interval across SAEs (shown as error bars in Figure 14). In addition, we performed paired bootstrap tests over SAEs for three contrasts: (i) Δ Token Confidence selection vs. SAE Base, and (ii) Δ Token Confidence selection vs. Output Score selection. Finally, we repeated the paired bootstrap test pooled over all 90 SAEs.

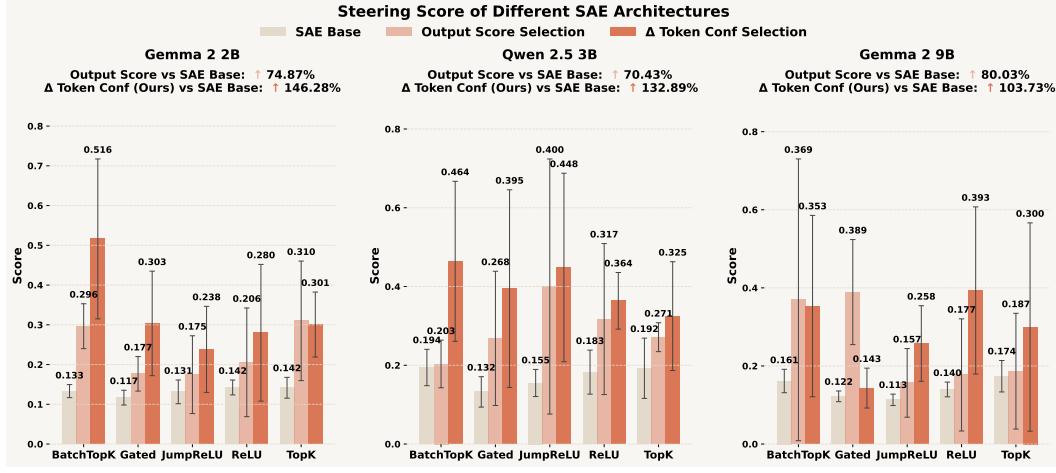


Figure 14: **Error-bar version of Figure 4.** Mean steering score and 95% confidence intervals across SAEs for the SAE Base, Output Score selection, and Δ Token Confidence selection, grouped by architecture and model. When aggregated across SAEs, there is a statistically significant difference between Δ Token Confidence and the baselines.

The error bars in Figure 14 show that Δ Token Confidence selection consistently lifts steering scores above the SAE Base across architectures and models.

In this experiment, steering scores are first computed for each of the 90 SAEs under three conditions: SAE Base (no feature selection), Output Score-based selection, and Δ TokenConf-based selection. We then form per-SAE paired differences and apply a paired bootstrap over SAEs to estimate the mean difference, its 95% confidence interval, and a two-sided p -value, yielding the global significance results summarized in Table 6.

1458
 1459 Table 6: Pooled paired-bootstrap comparison of steering scores across all 90 SAEs. For each con-
 1460 trast, we report the mean per-SAE difference in steering score, a 95% bootstrap confidence interval,
 1461 the two-sided p -value, and whether the effect is significant at the 5% level.

Comparison	Mean difference	95% bootstrap CI	Two-sided p	Significant at 5%?
Δ TokenConf – Base	0.1899	[0.1454, 0.2339]	≈ 0.0000	Yes
Δ TokenConf – Output Score	0.0786	[0.0261, 0.1305]	≈ 0.0032	Yes

1466
 1467 Concretely, we find that our selection method yields a statistically significant improvement in steer-
 1468 ing score both relative to the SAE Base and relative to the Output Score selection.

1471 K ANALYSIS OF GAP BETWEEN SAE INTERPRETABILITY AND UTILITY

1473 Regarding *why there is a gap*, our view is that the root cause lies in the training paradigm of SAEs.
 1474 Current SAE training is almost entirely reconstruction-centric: the objective is to accurately recon-
 1475 struct internal activations while enforcing sparsity, so as to obtain a more monosemantic basis in a
 1476 higher-dimensional space. This is precisely what makes SAE features interpretable, but it does not
 1477 directly optimize how useful those features are for steering. In other words, the training objective is
 1478 aligned with *reconstruction fidelity and monosemanticity*, not with *control over model outputs*.

1479 To make this concrete, we perform a pairwise Kendall– τ_b analysis over all 90 SAEs, relating inter-
 1480 pretability, steering, and two standard reconstruction-side metrics (both higher-is-better): (i) the **CE**
 1481 **loss score** and (ii) **explained variance**. Table 7 reports τ_b , bootstrap standard error, a 95% bootstrap
 1482 confidence interval, and a permutation p -value for the null $H_0: \tau_b = 0$ for each pair of metrics.

1485 Table 7: Pairwise Kendall– τ_b between interpretability, reconstruction metrics, and steering score
 1486 over all 90 SAEs. “CE loss score” and “explained variance” are standard post-hoc reconstruction
 1487 metrics (higher is better).

Pair	τ_b	SE	95% bootstrap CI	perm p ($H_0: \tau_b = 0$)
Interpretability – CE loss score	–0.433	0.056	[–0.539, –0.322]	0.0002
Interpretability – Explained variance	–0.405	0.059	[–0.521, –0.289]	0.0002
CE loss score – Steering	–0.243	0.067	[–0.377, –0.112]	0.0006
Explained variance – Steering	–0.195	0.066	[–0.323, –0.062]	0.0062

1494 As Table 7 shows, the reconstruction metrics exhibit only a *weak* association with steering ($|\tau_b| \approx$
 1495 0.2), but a substantially *stronger* association with interpretability ($|\tau_b| \approx 0.4$). In other words, the
 1496 reconstruction-focused training objective is much more predictive of which SAEs score well under
 1497 interpretability metrics than of which SAEs are actually good for steering, and is close to orthogonal
 1498 to steering utility. This supports our claim that the observed interpretability–utility gap is not an
 1499 artifact of our evaluation, but a structural consequence of the *current SAE training paradigm*, in
 1500 which reconstruction fidelity—rather than steering behavior—is the primary optimization target.

1503 L ABLATION ON THE TOP- k IN Δ TOKEN CONFIDENCE

1505 In our Δ token confidence selection, we measure how much a feature shifts the model’s confidence
 1506 over the top- k tokens in the output distribution. Unless otherwise noted, all main experiments use
 1507 $k = 1$, so that the score directly tracks the change in probability of the top-1 predicted token, which
 1508 typically dominates next-token behavior in standard generation settings.

1510 To verify that this choice is not artificially restrictive, we run an ablation over $k \in \{1, 3, 5, 10\}$ on
 1511 Gemma-2-2B and Qwen-2.5-3B. Table 8 reports the resulting steering scores (higher is better) after
 applying Δ token confidence selection.

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Table 8: Ablation over the top- k parameter in Δ token confidence selection.

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All tested values of k yield substantial gains over the SAE baseline (not shown), and $k = 1$ is in fact the best-performing setting on both models. We therefore adopt $k = 1$ as an empirically validated default in all main experiments.

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M EXAMPLES OF INTERPRETABLE-BUT-UNHELPFUL VS. HELPFUL-BUT-HARD-TO-EXPLAIN FEATURES

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To complement our quantitative analysis, we provide qualitative examples of features where interpretability and steering utility diverge. Type A features are *highly interpretable but not steerable* (high interpretability score, near-zero steering score), while Type B features are *moderately interpretable but strongly steerable*. All generations in Table 9 are produced with temperature 0 using the fixed prefix ‘‘I heard that’’ and intervening on a single feature.

1538

Table 9: Qualitative examples of features where interpretability and steering diverge. Type A: interpretable but not steerable (interpretability ≈ 1.0 , steering ≈ 0). Type B: moderately interpretable but clearly steerable. All generations use temperature 0 with the prefix “I heard that”.

Field	Gemma 2 2B (A)	Qwen 2.5 3B (A)	Gemma 2 2B (B)	Qwen 2.5 3B (B)
Type	A	A	B	B
Latent id	6168	12527	589	2392
Architecture	BatchTopK, $L_0=320$	BatchTopK, $L_0=160$	ReLU, $L_0=156$	JumpReLU, $L_0=166$
Interpretability	1.00	1.00	0.57	0.50
Steering score	0.00	0.00	1.32	1.02
Explanation	the symbol ‘pi’ in mathematical expressions	terms related to differentiation in mathematical contexts	words related to actions or processes in various contexts	numerical values related to comparisons
After steering	I heard that there are a lot of people who live with their own thoughts in their heads.	I heard that it is Math and Science, did you hear that it is Math and Science?	I heard that The Body Shop is planning to open in Central World, I used to go to Bangna branch.	I heard that a little bit larger than twice the actual, 2000 feet (2,500 meters) worth of this, which is a mile of it, three times it’s actual.

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