Group Fairness with Uncertainty in Sensitive Attributes

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Abstract

Learning a fair predictive model is crucial to mitigate biased decisions against minority groups in high-stakes applications. A common approach to learn such a model involves solving an optimization problem that maximizes the predictive power of the model under an appropriate group fairness constraint. However, in practice, sensitive attributes are often missing or noisy resulting in uncertainty. We demonstrate that solely enforcing fairness constraints on uncertain sensitive attributes can fall significantly short in achieving the level of fairness of models trained without uncertainty. To overcome this limitation, we propose a bootstrap-based algorithm that achieves the target level of fairness despite the uncertainty in sensitive attributes. The algorithm is guided by a Gaussian analysis for the independence notion of fairness where we propose a robust quadratically constrained quadratic problem to ensure a strict fairness guarantee with uncertain sensitive attributes. Our algorithm is applicable to both discrete and continuous sensitive attributes and is effective in real-world classification and regression tasks for various group fairness notions, e.g., independence and separation.

1. Introduction

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Achieving fairness in predictive modeling, whether in classification or regression tasks, is crucial to avoid discriminatory decisions against marginalized groups. Although various problem formulations exist for ensuring fairness in model training, a widely adopted approach is to formulate an optimization problem that maximizes the model's predictive power while satisfying a group fairness constraint (Kamishima et al., 2011; Zafar et al., 2017; Agarwal et al., 2018; Verma & Rubin, 2018; Gölz



Figure 1. Error vs fairness on Adult data for oracle and baseline models that enforce fairness using true and uncertain sensitive attributes, respectively. The baseline falls short of achieving the same range of fairness as the oracle.

et al., 2019; Mehrabi et al., 2021; Castelnovo et al., 2022). The notion of group fairness (Barocas et al., 2019) stipulates a certain (conditional) independence requirement involving the model prediction and the sensitive attribute. Then, the goal is to minimize the prediction loss while ensuring that the fairness loss, which measures the degree of group unfairness, i.e., the degree of violation of the (conditional) independence requirement, is less than a pre-defined tolerance level ϵ , i.e.,

min Prediction Loss s.t. Fairness Loss $\leq \epsilon$. (1)

Typically, it is assumed that the learner has access to true sensitive attributes for every sample in training, but in reality, labeled sensitive attributes are often missing or noisy. For instance, labeling sensitive attributes may require additional annotation of existing datasets for which such labels were not originally collected. Even if available, the sensitive attribute information can be uncertain due to various reasons, such as noisy or unreliable responses from survey participants due to fear of disclosure or discrimination (Krumpal, 2013). Moreover, privacy and legal regulations often limit the use of labeled sensitive attributes, such as race or gender, which are protected by laws such as the EU's General Data Protection Regulation or California's Consumer Privacy Act. In such cases, privatized sensitive attributes, which are obtained by adding noise, may be the only available option. In such scenarios, estimating the fairness loss in (1) using uncertain sensitive attributes, as if correct, can lead to a model that does

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Figure 2. Comparing fairness loss estimation methods: (a) Oracle uses true sensitive attributes $\mathcal{D}^{\text{oracle}}$; (b) Baseline uses available uncertain sensitive attributes $\mathcal{D}^{\text{uncertain}}$ as if correct; (c) Bootstrap-S constrains the optimization with additional fairness losses estimated using subsamples $\mathcal{D}_i^{\text{uncertain}}$, $\forall i \in [S]$. The horizontal bars illustrate these methods when uncertainty is due to missing sensitive attributes.

070 not accurately capture target fairness. Figure 1 shows 071 trade-off between prediction (measured by error rate) and 072 fairness (measured as violation of independence between predictions and sensitive attributes) obtained by varying ϵ 074 in (1) for Adult data (Lantz, 2019). The oracle (in red) 075 has access to true sensitive attributes, denoted by $\mathcal{D}^{\text{oracle}}$, enforces the fairness constraint: Fairness $Loss(\mathcal{D}^{oracle}) \leq \epsilon$, 077 and covers a wide range of fairness levels. In contrast, 078 the baseline (in orange) has access to a random < 1%of the sensitive attributes, denoted by $\mathcal{D}^{uncertain}$, enforces the fairness constraint: Fairness $Loss(\mathcal{D}^{uncertain}) \leq \epsilon$, but 081 is unable to achieve fairness below a threshold, i.e., the 082 baseline provides less control over attainable fairness 083 compared to the oracle.

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As a result, for high-stakes applications where violating a
fairness threshold incurs a significant cost, it is essential
to develop a method that can learn fair models despite
uncertainty in sensitive attributes.

Contributions. In this work, we propose a solution to fairlearning with uncertain sensitive attributes.

0921. We propose Bootstrap-S, an algorithm that uses a
bootstrap approach to impose S additional constraints to
the optimization in (1) for some parameter S. For $i \in [S]$,
constraint i requires Fairness Loss $(\mathcal{D}_i^{\text{uncertain}}) \leq \epsilon$, where
 $\mathcal{D}_i^{\text{uncertain}}$ is a collection of a fixed number of random
subsamples of the uncertain sensitive attributes $\mathcal{D}^{\text{uncertain}}$,
i.e., Bootstrap-S aims to

 $\min \operatorname{Prediction} \operatorname{Loss} \tag{2}$

s.t. Fairness
$$\text{Loss}(\mathcal{D}^{\text{uncertain}}) \leq \epsilon$$
 and,
Fairness $\text{Loss}(\mathcal{D}_i^{\text{uncertain}}) \leq \epsilon$ for all $i \in [S]$.

We illustrate Bootstrap-S in Figure 2 where it is contrasted with the oracle model that constrains the fairness loss estimated using true sensitive attributes $\mathcal{D}^{\text{oracle}}$ and the baseline model that constrains the fairness loss estimated using available uncertain sensitive attributes $\mathcal{D}^{uncertain}$ as if they are correct.

- 2. We analyze fair learning for Gaussian data with independence notion of fairness which serves as a motivation for Bootstrap-S. We begin by reducing a specific instance of this problem to a quadratically constrained quadratic problem (QCQP) when true sensitive attributes are available. Then, given uncertain sensitive attributes, we robustify the QCQP to provide a strict fairness guarantee. Notably, when uncertainty arises due to randomly missing sensitive attributes, the robust QCQP can achieve strict fairness without any performance loss, which we refer to as *free fairness*.
- 3. We showcase the practical effectiveness of Bootstrap-S in achieving fairness levels comparable to the oracle while maintaining high prediction performance (unlike the baseline) on a variety of real data, including classification and regression tasks, discrete and continuous sensitive attributes, and independence and separation notions of group fairness.

2. Problem Formulation

Suppose **x** represents *d*-dimensional input features defined on alphabet \mathcal{X} , while *y* and *e* denote 1-dimensional target and sensitive attribute defined on alphabets \mathcal{Y} and \mathcal{E} , respectively. Fair supervised learning seeks to find a predictor $f : \mathcal{X} \to \mathcal{Y}$ that (a) accurately estimates target variable for new input features and (b) avoids discrimination based on sensitive attribute. To achieve this, we are given (a) a loss function $\ell : \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}_+$, where $\ell(y, f(\mathbf{x}))$ measures the disagreement between target variable and its prediction, and (b) a fairness measure $\Phi : \mathcal{Y} \times \mathcal{Y} \times \mathcal{E} \to \mathbb{R}_+$, where $\Phi(y, f(\mathbf{x}), e)$ measures the level of discrimination of *f*. Given a fairness target $\epsilon \ge 0$ and a class of predictors \mathcal{F} , the goal of fair learning is to find an $f \in \mathcal{F}$ that minimizes expected loss ℓ , subject to fairness measure Φ being small:

$$f^* \in \operatorname*{arg\,min}_{f \in \mathcal{F}} \mathbb{E}[\ell(y, f(\mathbf{x}))] \quad \text{s.t.} \quad \Phi(y, f(\mathbf{x}), e) \le \epsilon.$$
 (3)

110 For ease of notation, hereon, we define $u \triangleq f(\mathbf{x})$.

111 Choice of fairness measure. To design a fairness measure 112 Φ , it is important to establish what is meant by a perfectly 113 fair predictor, i.e., $\epsilon = 0$ in (3). Typically, perfect fairness 114 is described in terms of statistical independence. We focus 115 on two commonly used fairness criteria: independence 116 and separation. The independence criterion, also called 117 demographic parity, demands $u \perp le$, meaning predictions 118 should not reveal any information about sensitive attributes. 119 The separation criterion, also known as equalized odds, 120 requires $u \perp | v |$, indicating predictions should not disclose 121 any information about sensitive attributes given the 122 knowledge of true target variables. 123

124 Achieving perfect fairness is not feasible when learning 125 a predictor from finite training samples (Agarwal et al., 126 2018). Instead, in practice, one often works with measures 127 of approximate fairness by choosing $\epsilon > 0$ in (3), and then 128 varying ϵ to find a balance between fairness and accuracy. 129 As perfect fairness measures assert that certain random 130 variables should be independent, a natural way to measure 131 approximate fairness is to use divergence that measures the 132 degree of independence between these variables. Recently, 133 χ^2 -divergence has emerged as an effective measure of 134 approximate fairness (Mary et al., 2019) and we adopt this 135 as our measure of the degree of independence, except in 136 cases where data is Gaussian, where we use a different analytically convenient divergence. For independence, 137 the measure is given by $\Phi(y, u, e) = \chi^2(p_{e,u} || p_e p_e)$ 138 139 where $p_{e,u}, p_e$, and p_u are marginals of (e, u), e, and u, respectively. Likewise, for separation $\Phi(y, u, e) =$ 140 141 $\mathbb{E}_{p_{y}}[\chi^{2}\left(p_{e,u|y} \| p_{e|y}p_{e|y}\right)] \text{ where } p_{e,u|y}, p_{e|y}, \text{ and } p_{u|y} \text{ are }$ 142 conditionals of (e, u), e, and u given y, respectively.

143 **Uncertain sensitive attributes.** Typically, N independent 144 and identically distributed (i.i.d.) samples of the 145 tuple (\mathbf{x}, y, e) are available, denoted by $\mathcal{D}^{(o)}$ ≜ $\{\mathbf{x}^{(i)}, \mathbf{y}^{(i)}, \mathbf{e}^{(i)}\}_{i \in [N]}$. Then, in the optimization in (3), 147 the objective is estimated using the subset $\mathcal{D}^{(p)} \triangleq$ 148 $\{\mathbf{x}^{(i)}, y^{(i)}\}_{i \in [N]}$ while the constraint is estimated using an 149 appropriate subset of $\mathcal{D}^{(o)}$ depending on the functional form 150 Φ . We denote these estimates by $\mathbb{E}_{\mathcal{D}^{(p)}}[\ell(y, f(\mathbf{x}))]$ and 151 $\Phi_{\mathcal{D}^{(o)}}(\mathbf{y}, f(\mathbf{x}), \mathbf{e})$, respectively, for brevity. We assume N 152 is sufficiently large and ignore any errors in these estimates 153 to focus on errors due to uncertainty in sensitive attributes. 154

155 When dealing with uncertain sensitive attributes, access 156 to $\mathcal{D}^{(o)}$ may not be possible. To account for such 157 uncertainty, we assume access to $\mathcal{D}^{(p)}$ as well as n < pN (potentially noisy) labeled sensitive attributes $\mathcal{D}^{(u)} \stackrel{\scriptscriptstyle{\frown}}{=}$ 158 159 $\{\mathbf{x}^{(i)}, y^{(i)}, \hat{e}^{(i)}\}_{i \in [n]}$. For $i \in [N]$, if $\hat{e}^{(i)} \neq e^{(i)}$, then 160 sensitive attribute $\hat{e}^{(i)}$ is noisy. Further, if n < N, then sensitive attributes $\{e^{(i)}\}_{i=n+1}^N$ are missing. Then, the goal 161 162 of fair learning with uncertain sensitive attributes is to solve 163

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the optimization in (3) with access to $\mathcal{D}^{(p)}$ and $\mathcal{D}^{(u)}$. While this is an intuitively appealing goal, simply computing the constraint in (3) with $\mathcal{D}^{(u)}$ may be sub-optimal as discussed in Section 1. In other words, a predictor *u* satisfying $\Phi_{\mathcal{D}^{(u)}}(y, u, e) \leq \epsilon$ may not necessarily satisfy $\Phi_{\mathcal{D}^{(o)}}(y, u, e) \leq \epsilon$.

In Appendix B, we consider the case where (x, y, e, u) is jointly Gaussian. Under this scenario, we fully characterize the optimization problem in (3) and guarantee strict fairness despite uncertainty in sensitive attributes. Then, we describe the motivation for our general-purpose algorithm by building on this analysis to develop.

3. A General-Purpose Algorithm

Now, we propose a generic algorithm, Bootstrap-S, to accounting for uncertainty in a non-parametric fashion following the *bootstrap* procedure (Efron, 1992). Specifically, given uncertain sensitive attribute data $\mathcal{D}^{(u)} = \{\mathbf{x}^{(i)}, y^{(i)}, \hat{\mathbf{e}}^{(i)}\}_{i \in [n]}$, a fairness measure $\Phi(y, u, e)$, and a parameter S: we draw uniformly S subsets $\mathcal{D}_1^{(u)}, \ldots, \mathcal{D}_S^{(u)}$ of some size $k \in [n]$ from $\mathcal{D}^{(u)}$ at random with replacement. Then, we estimate the fairness measure using each of these subsets as well as $\mathcal{D}^{(u)}$, and impose the collection of S constraints $\{\Phi_{\mathcal{D}_i^{(u)}}(y, u, e) \leq \epsilon\}_{i \in [S]}$ together with the constraint $\Phi_{\mathcal{D}^{(u)}}(y, u, e) \leq \epsilon$. In summary, we aim to solve:

$$\min_{u} \mathbb{E}_{\mathcal{D}^{(p)}} \left[\ell(y, u) \right] \text{ s.t } \Phi_{\mathcal{D}^{(u)}}(y, u, e) \le \epsilon \text{ and,} \tag{4}$$

$$\Phi_{\mathcal{D}^{(u)}}(\mathbf{y}, \mathbf{u}, \mathbf{e}) \leq \epsilon \text{ for all } i \in [S].$$

At a high level, the idea is similar to bootstrap confidence intervals (Wasserman, 2006) allowing construction of better uncertainty set with a larger number of subsamples S.

Notice that (4) is a constrained optimization problem, which is non-trivial to solve in practice, especially for neural network training. Typically, this problem is addressed by simply adding the fairness constraints as regularizers with hyperparameters to control the trade-off during optimization, i.e., min Prediction Loss + $\lambda \times$ Fairness Loss. However, the performance can be sub-optimal as it depends on the choice of λ . Instead, we follow the approach of Lee et al. (2019) by considering the Lagrangian dual of (4) and optimizing

Algorithm 1 Bootstrap-S
Input: $\mathcal{D}^{(p)}, \mathcal{D}^{(u)}$, subsample size $k \in [n]$, number of
subsamples $S \ge 1$.
for $i = 1$ to S do
Draw a subsample $\mathcal{D}_i^{(u)}$ of size k from $\mathcal{D}^{(u)}$ at random
with replacement.
end for
Then, solve the optimization problem in (5).

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Figure 3. Performance of Bootstrap-S and Baseline for independence (top row) and separation (bottom row). Bootstrap-S achieves much better fairness levels compared to Baseline throughout.

the resulting objective over the duality variables, i.e.,

$$\min_{u} \max_{\lambda, \lambda_{1}, \cdots, \lambda_{S}} \mathbb{E}_{\mathcal{D}^{(p)}} \left[\ell(y, u) \right] + \lambda \left(\Phi_{\mathcal{D}^{(u)}}(y, u, e) - \epsilon \right) \quad (5) \\
+ \sum_{i \in [S]} \lambda_{i} \left(\Phi_{\mathcal{D}_{i}^{(u)}}(y, u, e) - \epsilon \right)$$

We summarize Bootstrap-S in Algorithm 1.

4. Empirical Evaluation

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We test Bootstrap-S on real-world classification and regression tasks for group fairness notions of independence and separation using 3 datasets: Adult, Crime, & Insurance. We provide an overview of these in Table 1 with detailed descriptions and pre-processing steps in Appendix I.2.

Table 1. Overview of datasets.

Dataset	Task	Sensitive Attribute
Adult	Classification	Sex (binary)
Crime	Regression	Race (continuous)
Insurance	Regression	Sex (binary)

Training. We train a two-layered neural network and use the log loss for classification, and MSE for regression. We use the χ^2 -divergence to impose the independence (conditional independence) for independence (separation). For continuous sensitive attributes, we induce uncertainty in every sensitive attribute by adding independent $\mathcal{N}(0, \sigma^2)$ noise ($\sigma = 0.5$ for Crime). For binary sensitive attributes, we induce uncertainty by keeping only *n* out of *N* sensitive attributes (n = 100 for Adult and n = 10 Insurance). For Bootstrap-S, we set S = 5. Given a fairness target ϵ , we train a model over 50 independent trials of random missingness (for Adult and Insurance) or random noise (for Crime), and report the average performance (the error bars are too small to see). We sweep over 500 different ϵ from 0.001 to 0.5, and plot the prediction-fairness trade-off frontier by using a simple moving average over 5 entries. We provide more details and experiments in Appendix I.2

Evaluation. On a held-out test set, we report predictive power using error rate (lower is better) for classification and MSE (lower is better) for regression. We evaluate (a) independence (lower is better) using demographic parity, i.e., $|P(\hat{y}|e = 1) - P(\hat{y}|e = 0)|$ for classification and χ^2 -divergence for regression, and (b) separation (lower is better) using equal opportunity, i.e., $|P(\hat{y}|e = 1, y = 1) - P(\hat{y}|e=0, y=1)|$ for classification and χ^2 -divergence for regression.

Results. We compare with Baseline which does not have the additional constraints of Bootstrap-S, i.e., it solves for (5) with $\lambda_1, \dots, \lambda_S$ fixed to 0. For reference, we also compare with Oracle that has access to all the true sensitive attributes. Figure 3 show that Baseline and Bootstrap-S exhibit a concentration of fairness levels near the extreme values of ϵ with Bootstrap-S being more noisy than Baseline. However, Bootstrap-S achieves significantly smaller fairness levels compared to Baseline. In fact, in most cases, Bootstrap-S achieves fairness levels that are comparable to Oracle while maintaining a relatively high predictive power.

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0 A. Related Work

Group fairness in machine learning. Various metrics and criteria have been proposed to ensure group fairness in machine learning (Verma & Rubin, 2018; Mehrabi et al., 2021; Castelnovo et al., 2022; Shah et al., 2022), but many of these criteria are mutually exclusive in non-trivial cases (Gölz et al., 2019). For example, the independence and the separation criteria cannot both be satisfied simultaneously. Different approaches exist to enforce these criteria, mainly falling into one of three categories: (*a*) pre-processing methods (Zemel et al., 2013; Calmon et al., 2017), (*b*) post-processing methods (Hardt et al., 2016; Pleiss et al., 2017), and (*c*) in-processing methods (Kamishima et al., 2011; Zafar et al., 2017; Agarwal et al., 2018). In this work, we consider independence and separation, using an in-processing approach where the objective function accounts for both accuracy and fairness.

Fairness without certain sensitive attributes. The growing literature on fairness in the absence of true sensitive attributes
 can be broadly categorized into the following three groups.

A. Perturbed sensitive attribute. Several approaches have been proposed to handle perturbations in sensitive attributes. For
example, in-processing methods for fair classification have been developed by (Lamy et al., 2019) and (Celis et al., 2021a)
to deal with noisy group labels, while (Awasthi et al., 2020) investigated the performance of a post-processing algorithm
proposed by (Hardt et al., 2016) with noisy sensitive labels. Additionally, (Mozannar et al., 2020) and (Celis et al., 2021b)
explored achieving group fairness with adversarially perturbed and differentially private data, respectively.

B. Proxy variables. (Gupta et al., 2018), (Chen et al., 2019), and (Kallus et al., 2022) proposed methods to achieve group
fairness when proxy variables are available as substitutes for the sensitive attribute (e.g., zip code as a proxy of race).
However, the effectiveness of these methods may be reduced if the correlation between the sensitive attribute and the proxy
variables is weak. (Jung et al., 2022) proposed a semi-supervised learning approach to generate proxy pseudo-labels for
partially observed sensitive attributes. While these proxy-based methods can be useful, they risk perpetuating biases.

C. No sensitive attribute. (Hashimoto et al., 2018) and (Lahoti et al., 2020) proposed methods for achieving fairness without relying on a labeled sensitive attribute, utilizing distributionally robust optimization to improve the performance of the worst-case risk for all distributions close to the empirical distribution. They aim to achieve Rawlsian max-min fairness, but their notion of fairness is not defined by the population distribution, which sets it apart from our focus on group fairness. Additionally, it may not be straightforward to combine their methods with existing fair-training methods, while our method can be generally applied to any fair-training method.

Our work is most closely related to (Wang et al., 2020), who also focus on achieving a strict group fairness guarantee given uncertain sensitive attributes. However, they focus on classification problems with discrete sensitive attributes. In contrast, our approach is widely applicable, including both regression and classification, as well as both discrete and continuous sensitive attributes.

B. Gaussian setting

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For the ease of exposition, we consider zero-mean Gaussian variables, and assume that the marginal distribution $p_{x,y}$ is known or can be learned from $\mathcal{D}^{(p)}$.n We think of *u* as a representation of the features and let the predictor be $\mathbb{E}[y|u]$. Naturally, the loss function ℓ is chosen to be the mean squared loss as $\mathcal{Y} = \mathbb{R}$. We focus on the independence criterion of fairness and measure the degree of independence between *u* and *e* using the notion of \overline{D} -divergence, a second-order approximation of Kullback–Leibler divergence, introduced by (Huang et al., 2019).

Definition 1 (\overline{D} -divergence). The \overline{D} -divergence between zero-mean Gaussian random vectors $\mathbf{v} \sim p_{\mathbf{v}} = \mathcal{N}(\mathbf{0}, \Sigma_v)$ and $\mathbf{w} \sim p_{\mathbf{w}} = \mathcal{N}(\mathbf{0}, \Sigma_w)$, with $\|\cdot\|_{\mathrm{F}}$ denoting the Frobenius norm, is given by

$$\overline{D}(p_{\mathbf{v}} \| p_{\mathbf{w}}) \triangleq \frac{1}{2} \| \mathbf{\Sigma}_{w}^{-1/2} (\mathbf{\Sigma}_{v} - \mathbf{\Sigma}_{w}) \mathbf{\Sigma}_{w}^{-1/2} \|_{\mathrm{F}}^{2}, \tag{6}$$

For these choices, the optimization in (3) reduces to learning a Gaussian variable u such that

$$u^* \in \operatorname*{arg\,min}_{u} \mathbb{E}\left[(y - \mathbb{E}[y|u])^2\right] \quad \text{s.t.} \quad \overline{D}(p_{e,u}||p_e p_u) \le \epsilon.$$
(7)

Next, we reformulate (7) into a quadratically constrained quadratic program (QCQP) by utilizing the notion of canonical correlation matrices (CCMs) defined by (Huang et al., 2019).

Definition 2 (Canonical correlation matrix). The canonical correlation matrix (CCM) between zero-mean jointly Gaussian random vectors $\mathbf{v} \sim \mathcal{N}(\mathbf{0}, \Sigma_v)$ and $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \Sigma_w)$ is given by $\mathbf{b}_{vw} \triangleq \Sigma_{vv}^{-1/2} \Sigma_{vw} \Sigma_{ww}^{-1/2}$, where Σ_{vw} is the cross-covariance matrix between \mathbf{v} and \mathbf{w} .

The \overline{D} -divergence is conveniently represented by these CCMs. We now formally state the equivalence between (7) and a QCQP that uses CCMs. We prove this in Appendix C.1 by drawing connections to the information bottleneck principle (Bu et al., 2021).

Theorem B.1 (Gaussian Fair Learning \iff QCQP). The optimization problem in (7) is equivalent to

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \quad s.t. \quad \left\langle \mathbf{a}, \mathbf{b}_{ex} \right\rangle^2 \le \varepsilon, \tag{8}$$

where $\mathcal{B}(0,1)$ denotes an ℓ_2 ball centered at 0 with radius I, $\langle \cdot, \cdot \rangle$ denotes the inner product, and **a** plays the role of **b**_{ux}.

We note that **a** in (8) has the same dimension as \mathbf{x} , i.e., *d*. The following result, with a proof in Appendix C.2, shows that any *d*-dimensional QCQP in (8) can be mapped to a 2-dimensional QCQP.

Proposition 1 (d = 2 suffices for QCQP). The optimal solution \mathbf{a}^* of the QCQP in (8) lies in the subspace spanned by the vectors \mathbf{b}_{yx} and \mathbf{b}_{ex} .

In Appendix C.3, we characterize the optimal \mathbf{a}^* in (8) as a function of \mathbf{b}_{yx} , \mathbf{b}_{ex} , and ε for d = 2. Theorem B.1 demonstrates that considering the uncertainty in the canonical correlation matrix b_{ex} is sufficient to capture the uncertainty in sensitive attributes, which we will now explore in detail.

B.1. Theoretical Results

In this section, we provide a characterization of fair learning for Gaussian data given some uncertainty in sensitive attributes. Specifically, we study how to *robustify* the QCQP in (8) to ensure strict fairness guarantee with high probability, as well as how this robustification affects the optimal objective.

Let $\hat{\mathbf{b}}_{ex}$ be an estimate of \mathbf{b}_{ex} , say obtained from $\mathcal{D}^{(u)}$, such that $\|\mathbf{b}_{ex} - \hat{\mathbf{b}}_{ex}\|_2 \leq \tau$ (with probability $1 - \delta$), for some $\tau \geq 0,^1$ and denoted by $\mathbf{b}_{ex} \in \mathcal{B}(\hat{\mathbf{b}}_{ex}, \tau)$. To achieve fairness as in (8) with probability $1 - \delta$, in the worst case, $\langle \mathbf{a}, \mathbf{b} \rangle^2 \leq \varepsilon$ should hold for all $\mathbf{b} \in \mathcal{B}(\hat{\mathbf{b}}_{ex}, \tau)$. Then, the following robust optimization maximizes the desired objective while achieving fairness as in (8) (with probability $1 - \delta$) without the precise knowledge of \mathbf{b}_{ex} :

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)} \langle \mathbf{a}, \mathbf{b}_{yx} \rangle^2 \quad \text{s.t.} \quad \langle \mathbf{a}, \mathbf{b} \rangle^2 \le \varepsilon, \ \forall \mathbf{b}\in\mathcal{B}(\widehat{\mathbf{b}}_{ex}, \tau).$$
(9)

In the following proposition, we show that any *d*-dimensional robust QCQP in (9) can be mapped to a 2-dimensional QCQP. See Appendix D for a proof.

Proposition 2 (d = 2 suffices for robust QCQP). The optimal solution \mathbf{a}^* of the robust QCQP in (9) lies in the subspace spanned by the vectors \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$.

Now, to characterize the solution of the robust QCQP in (9), we focus on d = 2 and work with polar coordinates. Further, to analyze the corresponding feasible space, we relax the uncertainty space $\mathcal{B}(\hat{\mathbf{b}}_{ex},\tau)$ from a ball to an annular sector. Formally, we let $\hat{\mathbf{b}}_{ex} \triangleq \hat{r}_e(\cos\hat{\theta}_e, \sin\hat{\theta}_e)$ be the estimate of \mathbf{b}_{ex} such that $|r_e - \hat{r}_e| \le \Delta$ and $|\theta_e - \hat{\theta}_e| \le \phi$ with probability $1 - \delta$ where $\Delta \triangleq \tau \ge 0$ and $\phi \triangleq \sin^{-1}(\tau/\hat{\mathbf{b}}_{ex}) \in [0, \pi/2]$. In other words, given $\hat{r}_e, \hat{\theta}_e, \Delta$, and ϕ , with probability $1 - \delta$,

$$\mathbf{b}_{ex} \in \mathcal{A}(\Delta, \phi) \triangleq \{ \mathbf{b} = r(\cos\theta, \sin\theta) \colon |r - \hat{r}_e| \le \Delta \text{ and } |\theta - \hat{\theta}_e| \le \phi \},\tag{10}$$

i.e., $\mathcal{A}(\Delta, \phi)(\supset \mathcal{B}(\widehat{\mathbf{b}}_{ex}, \tau))$ denotes the smallest annular sector around $\widehat{\mathbf{b}}_{ex}$ capturing our uncertainty in knowing \mathbf{b}_{ex} (see Figure 4 where $\mathcal{A}(\Delta, \phi)$ is the shown in orange). Now, to achieve fairness as in (8) (with probability $1 - \delta$), we constrain the robust QCQP in (9) as follows:

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)} \langle \mathbf{a}, \mathbf{b}_{yx} \rangle^2 \quad \text{s.t.} \quad \langle \mathbf{a}, \mathbf{b} \rangle^2 \le \varepsilon, \ \forall \mathbf{b} \in \mathcal{A}(\Delta, \phi).$$
(11)

¹For ease of the exposition, we assume $\tau \leq \|\mathbf{b}_{ex}\|_2$.



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(a) Feasible space when $\sqrt{\varepsilon} > (\hat{r}_e + \Delta) \sin \phi$. (b) Feasible space when $\sqrt{\varepsilon} < (\hat{r}_e + \Delta) \sin \phi$.

Figure 4. Visualizing the feasible space of the robust QCQP in Theorem B.2, i.e., (12), for $\varepsilon = 0.9$, $\hat{r}_e = 1.6$, and $\hat{\theta}_e = 0$. We set $\Delta = 0.2$ and $\phi = \pi/12$ for panel (a), and $\Delta = 0.4$ and $\phi = \pi/6$ for panel (b). Each point is shown in polar coordinates, i.e., a point (r, θ) denotes ($r \cos \theta, r \sin \theta$). The annular sector $\mathcal{A}(\Delta, \phi)$ is the region enclosed by dashed lines, dashed arc, and solid arc in orange. The arc $\overline{\mathcal{A}}(\Delta, \phi)$ is the solid arc in orange. The shaded region is the feasible space. The points $\mathbf{b}_{ex}^{(1)}$, $\mathbf{b}_{ex}^{(2)}$, and $\mathbf{b}_{ex}^{(3)}$ from Theorem B.3 are shown in magenta, blue and green, respectively.

As we show below, the constraint in (11) is equivalent to ensuring $\langle \mathbf{a}, \mathbf{b} \rangle^2 \leq \varepsilon$ for all $\mathbf{b} \in \overline{\mathcal{A}}(\Delta, \phi)$ where $\overline{\mathcal{A}}(\Delta, \phi)$ is the arc on the boundary of the angular sector $\mathcal{A}(\Delta, \phi)$ with maximum radius (shown in solid orange in Figure 4). See Appendix E for a proof. Further, in Appendix E.1, we characterize the optimal \mathbf{a} in (11) as a function of $\hat{\mathbf{b}}_{ex}$, \mathbf{b}_{yx} , Δ , ϕ , and ε . The same **B** 2 (B short OCOP with infinite constraints). Let $\overline{\mathcal{A}}(\Delta, \phi) \triangleq (\mathbf{b} + \mathbf{b}) = (\widehat{\omega} + \Delta)(\cos \theta, \sin \theta)$, and $|\theta| = \widehat{\mathcal{A}} + \varepsilon$.

Theorem B.2 (Robust QCQP with infinite constraints). Let $\overline{\mathcal{A}}(\Delta, \phi) \triangleq \{\mathbf{b} : \mathbf{b} = (\widehat{r}_e + \Delta)(\cos \theta, \sin \theta) \text{ and } |\theta - \widehat{\theta}_e| \le \phi\}$ be the arc on the boundary of $\mathcal{A}(\Delta, \phi)$ with maximum radius. Then, the robust QCQP in (11) is equivalent to

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)} \langle \mathbf{a}, \mathbf{b}_{yx} \rangle^2 \quad s.t. \quad \left\langle \mathbf{a}, \mathbf{b} \right\rangle^2 \le \varepsilon, \ \forall \mathbf{b} \in \overline{\mathcal{A}}(\Delta, \phi).$$
(12)

There is a phase transition in the nature of the feasible space of the robust QCQP in (12). Figure 4(a) illustrates the space if $\sqrt{\varepsilon} \ge (\hat{r}_e + \Delta) \sin \phi$, and Figure 4(b) illustrates the space if $\sqrt{\varepsilon} \le (\hat{r}_e + \Delta) \sin \phi$.

While Theorem B.2 simplifies the optimization in (11), the resulting optimization in (12) still has infinite constraints. Below, we provide an approximation to the feasible space in (12) such that it has finitely many constraints. See Appendix F for a proof. We note that alternative approximations are possible.

Theorem B.3 (Robust QCQP with 3 constraints). Let $\mathbf{b}_{ex}^{(1)} = \frac{(\hat{r}_e + \Delta)}{\cos \phi} (\cos \hat{\theta}_e, \sin \hat{\theta}_e), \ \mathbf{b}_{ex}^{(2)} = (\hat{r}_e + \Delta)(\cos(\hat{\theta}_e + \phi), \sin(\hat{\theta}_e + \phi)), \ and \ \mathbf{b}_{ex}^{(3)} = (\hat{r}_e + \Delta)(\cos(\hat{\theta}_e - \phi), \sin(\hat{\theta}_e - \phi)).$ Then, the feasible space of the optimization below is a subset of the feasible space of the optimization in (12):

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \quad s.t \quad \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(i)} \right\rangle^2 \le \varepsilon \text{ for all } i \in [3].$$
(13)

See Appendix F.1 for a characterization of the optimal **a** in (13) as a function of $\hat{\mathbf{b}}_{ex}$, \mathbf{b}_{yx} , Δ , ϕ , and ε . We visualize $\mathbf{b}_{ex}^{(1)}$, $\mathbf{b}_{ex}^{(2)}$, and $\mathbf{b}_{ex}^{(3)}$ in Figure 4 (in magenta, blue, and green, respectively), and note that $\mathbf{b}_{ex}^{(1)}$ approximates the effect of the points in-between $\mathbf{b}_{ex}^{(2)}$ and $\mathbf{b}_{ex}^{(3)}$ on $\mathcal{A}(\Delta, \phi)$.

Uncertainty due to randomly missing sensitive attributes. Now, we focus on understanding how the optimal objective of the robust QCQP in (13) changes when the uncertainty set $\mathcal{A}(\Delta, \phi)$ changes. For concreteness, we consider the case where uncertainty only stems from sensitive attributes missing at random, and the uncertainty can be improved by collecting more labeled sensitive attributes. Then, we analyze the power of each new labeled sensitive attribute by characterizing the exact difference in the optimal objectives of the QCQP in (8) and the robust QCQP in (13). To enable this, we show (in Appendix H) that the optimal objective in (13) either monotonically increases or coincides with the optimal objective in (8) whenever the uncertainty set monotonically decreases with *n*.

The response of our algorithm to collecting more labeled sensitive attributes can be classified into 3 categories (see the formal result in Appendix G): (a) Any uncertainty hurts: Here, the optimal objective of the robust QCQP in (13) matches

robust QCQP in (13) matches the optimal objective in (8) without collecting any extra labeled sensitive attributes. We state

499 this informally below (see Lemma 4 for a formal statement).

Corollary 1 (Free fairness). There exist problem instances of the robust QCQP in (13) where the uncertainty incurs no
 performance loss while achieving a strict fairness guarantee, without requiring additional labeled sensitive attributes.

Solving the robust QCQP in (13) for d = 2 is straightforward using a standard convex optimization solver. However, as the dimensionality of x increases, solving the problem becomes more challenging. Furthermore, real datasets may not be well-modeled by Gaussian distributions, and thus, the proposed robust QCQP may not be directly deployable.

507 B.2. Motivation for Bootstrap-S

In Section 3, we leverage our theoretical analysis to propose a generic algorithm that handles high-dimensional features and non-Gaussian data while accounting for uncertainty. The core idea of the robust QCQP (in Theorem B.3) is to construct an uncertainty set around the estimated canonical correlation matrix $\hat{\mathbf{b}}_{ex}$ by imposing additional constraints. This approach effectively addresses the unknown nature of the true \mathbf{b}_{ex} . Another perspective is to view the robust QCQP in (13) as

$$\max_{\mathbf{a}\in\mathcal{B}(0,1)}\left\langle \mathbf{a},\mathbf{b}_{yx}\right\rangle^{2} \quad \text{s.t} \quad \left\langle \mathbf{a},\widehat{\mathbf{b}}_{ex}\right\rangle^{2} \leq \varepsilon \quad \text{and} \quad \left\langle \mathbf{a},\mathbf{b}_{ex}^{(i)}\right\rangle^{2} \leq \varepsilon \quad \text{for all } i \in [3], \tag{14}$$

515 516 where the constraint $\langle \mathbf{a}, \hat{\mathbf{b}}_{ex} \rangle^2 \le \varepsilon$ becomes redundant in the presence of the constraint $\langle \mathbf{a}, \mathbf{b}_{ex}^{(1)} \rangle^2 \le \varepsilon$, and $\{\mathbf{b}_{ex}^{(i)}\}_{i \in [3]}$ can 517 be viewed as multiple estimates of $\hat{\mathbf{b}}_{ex}$. For general non-Gaussian data, we use a similar idea in (4).

519 **B.3. Synthetic data**

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In this section, we show the efficacy of the robust QCQP in achieving strict fairness on synthetic Gaussian data and
 demonstrate that Bootstrap-S serves as a good approximation for the robust QCQP.

523 We generate synthetic data using two Gaussian distributions with zero mean, d = 2, and covariances matrices (a) Σ_2^{gen} and 524 (b) Σ_2^{fair} defined in Appendix I.1. We present our results when uncertainty is due to randomly missing sensitive attributes. In 525 Appendix I.1, we present results on (i) another choice of covariance, (ii) d = 3, and (iii) uncertainty in sensitive attribute 526 due to noise. We estimate $\hat{\mathbf{b}}_{ex}$ using *n* samples of (\mathbf{x}, e) for various choices of *n*. Then, we compare the robust QCQP 527 in (13) and Bootstrap-S applied to the QCQP in (8) (for various S) against Baseline, which solves the QCQP in 528 (8) using $\hat{\mathbf{b}}_{ex}$. We solve the resulting optimization problems using the CVXPY library (Diamond & Boyd, 2016). The covariances Σ_2^{gen} an Σ_2^{fair} are designed to demonstrate the general behavior (where uncertainty hurts) and the free-fairness 529 530 behavior (Corollary 1), respectively. 531

The results, averaged over 1000 random trials, are shown in Figure 5 (the error bars are too small to see). We observe 532 533 that robust QCQP always ensures no fairness violations. Additionally, the performance (in terms of average MSE) of robust QCQP monotonically improves as n increases, as stated in Appendix H. More importantly, in Figure 5 (b), robust 534 QCQP does not incur any significant loss in the performance, and thus demonstrates the free-fairness phenomenon in 535 Corollary 1, say, $n \approx 350$ onward. We also see that Bootstrap-S well approximates the performance of robust QCQP 536 and outperforms Baseline in terms of fairness violations. As alluded to earlier, Bootstrap-S achieves a better fairness 537 criterion as we increase the number of subsamples S by forming a more accurate uncertainty set. However, the benefit of 538 larger S comes with an increased computation. 539

541 C. Proof of Theorem B.1 and Proposition 1 542

543 In this section, we prove Theorem B.1 and Proposition 1 as well as provide a characterization of the optimal \mathbf{a}^{\star} in (8).

545 C.1. Proof of Theorem B.1: Gaussian Fair Learning \iff QCQP

546 547 First, we show that optimization in (7) is equivalent to the optimization below:

$$\arg\max_{u} \overline{D}(p_{y,u} \| p_{y} p_{u}) \quad \text{s.t.} \quad \overline{D}(p_{e,u} \| p_{e} p_{u}) \le \epsilon.$$
(15)

Group Fairness with Uncertainty in Sensitive Attributes



Figure 5. The performance of robust QCQP in (13), Bootstrap-S with $S \in \{3, 9, 27\}$, and Baseline for d = 2 and various Σ^* . In the left column, we plot the fraction of violations of the true fairness constraint $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$ vs. n; in the middle column, we plot average MSE vs. n; in the right column, we plot the histogram of the value of $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2$ over 1,000 trials for n = 250.

To that end, from the definition of conditional variance, we have

$$\mathbb{E}[(y - \mathbb{E}[y|u])^2] = \mathbb{E}[\mathbb{V}\mathrm{ar}[y|u]].$$
(16)

Now, for a joint Gaussian vector $(y, u) \in \mathbb{R}^2$ with covariance Σ , we have

$$\mathbb{E}\left[\mathbb{V}\mathrm{ar}[y|u]\right] \stackrel{(a)}{=} \mathbb{E}\left[\boldsymbol{\Sigma}_{yy} - \boldsymbol{\Sigma}_{yu}\boldsymbol{\Sigma}_{uu}^{-1}\boldsymbol{\Sigma}_{uy}\right] = \boldsymbol{\Sigma}_{yy} - \boldsymbol{\Sigma}_{yu}\boldsymbol{\Sigma}_{uu}^{-1}\boldsymbol{\Sigma}_{uy},\tag{17}$$

where (a) follows from Schur's complement. Since Σ_{yy} is a constant w.r.t. u, we can write

$$\arg\min_{u} \mathbb{E}\left[\mathbb{V}\mathrm{ar}[y|u]\right] = \arg\min_{u} -\Sigma_{yu} \Sigma_{uu}^{-1} \Sigma_{uy} = \arg\max_{u} \Sigma_{yu} \Sigma_{uu}^{-1} \Sigma_{uy}.$$
(18)

Now, from Huang et al. (2019, Lemma. 68), we have

$$\overline{D}(p_{y,u}||p_yp_u) = \|\boldsymbol{\Sigma}_{yy}^{-1/2}\boldsymbol{\Sigma}_{yu}\boldsymbol{\Sigma}_{uu}^{-1/2}\|_{\mathrm{F}}^2 \stackrel{(a)}{=} \boldsymbol{\Sigma}_{yy}^{-1/2}\boldsymbol{\Sigma}_{yu}\boldsymbol{\Sigma}_{uu}^{-1}\boldsymbol{\Sigma}_{uy}\boldsymbol{\Sigma}_{yy}^{-1/2},$$
(19)

where $\| \cdot \|_{\mathrm{F}}$ denotes the Frobenius norm, and (a) follows because $\Sigma_{yy}^{-1/2} \Sigma_{yu} \Sigma_{uu}^{-1} \Sigma_{uy} \Sigma_{yy}^{-1/2} \in \mathbb{R}$. As before, since Σ_{yy} is a constant w.r.t. u, we can write

$$\arg\max_{u} \overline{D}(p_{y,u} \| p_{y} p_{u}) = \arg\max_{u} \Sigma_{yu} \Sigma_{uu}^{-1} \Sigma_{uy}.$$
(20)

Then, (15) follows by combining (16), (18), and (20).

Next, from Bu et al. (2021, Theorem. 1), the optimization problem in (15) is equivalent to the following semi-definite program (SDP) using the notion of canonical correlation matrices (Definition 2):

$$\max_{\mathbf{A} \in \mathbb{A}^d} \operatorname{Tr}(\mathbf{b}_{yx}^\top \mathbf{b}_{yx} \mathbf{A}) \quad \text{s.t.} \quad \operatorname{Tr}(\mathbf{b}_{ex}^\top \mathbf{b}_{ex} \mathbf{A}) \le \varepsilon \text{ and } 0 \preceq \mathbf{A} \preceq \mathbf{I},$$
(21)

where $\text{Tr}(\cdot)$ denote the trace of a matrix, \mathbb{A}^d is the space of $d \times d$ symmetric matrices, and **A** is of the form $\mathbf{b}_{xu}\mathbf{b}_{xu}^{\top}$. Finally, we show that the SDP in (21) is equivalent to the QCQP in (8). First, we have

$$\operatorname{Tr}(\mathbf{b}_{yx}^{\top}\mathbf{b}_{yx}\mathbf{A}) \stackrel{(a)}{=} \operatorname{Tr}(\mathbf{b}_{yx}^{\top}\mathbf{b}_{yx}\mathbf{b}_{xu}\mathbf{b}_{xu}^{\top}) \stackrel{(b)}{=} \operatorname{Tr}(\mathbf{b}_{xu}^{\top}\mathbf{b}_{yx}^{\top}\mathbf{b}_{yx}\mathbf{b}_{xu}) = \operatorname{Tr}((\mathbf{b}_{yx}\mathbf{b}_{xu})^{\top}\mathbf{b}_{yx}\mathbf{b}_{xu}) \stackrel{(c)}{=} \langle \mathbf{b}_{yx}, \mathbf{b}_{xu}^{\top} \rangle^{2}, \quad (22)$$

where (a) follows because $\mathbf{A} = \mathbf{b}_{xu} \mathbf{b}_{xu}^{\top}$, (b) follows from the cyclic property of trace, and (c) follows because $\langle \mathbf{b}_{yx}, \mathbf{b}_{xu}^{\top} \rangle = \mathbf{b}_{yx} \mathbf{b}_{xu} \in \mathbb{R}^{1 \times 1}$. Similarly, we have

$$\operatorname{Ir}(\mathbf{b}_{ex}^{\top}\mathbf{b}_{ex}\mathbf{A}) = \left\langle \mathbf{b}_{ex}, \mathbf{b}_{xu}^{\top} \right\rangle^{2}.$$
(23)

Finally, noting that A is a rank 1 matrix, we have

$$0 \leq \mathbf{A} \leq \mathbf{I} \iff \mathbf{b}_{xu} \in \mathcal{B}(0,1).$$
⁽²⁴⁾

Putting together (21) to (24) completes the proof.

C.2. Proof of Proposition 1: d = 2 suffices for QCQP

617 It suffices to show that the projection of any $\mathbf{a} \in \mathcal{B}(0, 1)$, satisfying the constraint $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$, onto the subspace spanned 618 by \mathbf{b}_{yx} and \mathbf{b}_{ex} preserves the value of the objective and continues to satisfy the constraint. Fix some \mathbf{a} in the feasible space. 619 Let $\operatorname{proj}(\mathbf{a})$ denote the projection of \mathbf{a} on the subspace spanned by \mathbf{b}_{yx} and \mathbf{b}_{ex} .

The value of the objective is preserved. It suffices to show $\langle \text{proj}(\mathbf{a}), \mathbf{b}_{yx} \rangle^2 = \langle \mathbf{a}, \mathbf{b}_{yx} \rangle^2$. Consider an orthonormal basis $\{\mathbf{n}_y, \mathbf{n}_e\} \triangleq \{\frac{\mathbf{b}_{yx}}{\|\mathbf{b}_{yx}\|_2}, \mathbf{n}_e\}$ spanned by \mathbf{b}_{yx} and \mathbf{b}_{ex} where $\mathbf{n}_y = \frac{\mathbf{b}_{yx}}{\|\mathbf{b}_{yx}\|_2}$ and \mathbf{n}_e is chosen to be orthogonal to \mathbf{n}_y . Then,

$$\operatorname{proj}(\mathbf{a}) = \langle \mathbf{a}, \mathbf{n}_y \rangle \mathbf{n}_y + \langle \mathbf{a}, \mathbf{n}_e \rangle \mathbf{n}_e.$$
(25)

626 Using the definition of \mathbf{n}_y and the orthogonality between \mathbf{n}_e and \mathbf{b}_{yx} , we have

$$\langle \operatorname{proj}(\mathbf{a}), \mathbf{b}_{yx} \rangle^2 = \langle \mathbf{a}, \mathbf{n}_y \rangle^2 \|\mathbf{b}_{yx}\|_2^2.$$
 (26)

630 Similarly, using the definition of \mathbf{n}_y , we have $\langle \mathbf{a}, \mathbf{b}_{yx} \rangle^2 = \langle \mathbf{a}, \mathbf{n}_y \rangle^2 ||\mathbf{b}_{yx}||_2^2$.

The constraint is satisfied. It suffices to show $\operatorname{proj}(\mathbf{a}) \in \mathcal{B}(0,1)$ and $\langle \operatorname{proj}(\mathbf{a}), \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$. It is easy to see $\operatorname{proj}(\mathbf{a}) \in \mathcal{B}(0,1)$ because $\|\operatorname{proj}(\mathbf{a})\|_2 \leq \|\mathbf{a}\|_2$. Now, consider a different set of orthonormal basis $\{\mathbf{n}'_e, \mathbf{n}'_y\} \triangleq \{\frac{\mathbf{b}_{ex}}{\|\mathbf{b}_{ex}\|_2}, \mathbf{n}'_y\}$ spanned by \mathbf{b}_{ex} and \mathbf{b}_{yx} where $\mathbf{n}'_e = \frac{\mathbf{b}_{ex}}{\|\mathbf{b}_{ex}\|_2}$ and \mathbf{n}'_y is chosen to be orthogonal to \mathbf{n}'_e . Then,

$$\operatorname{proj}(\mathbf{a}) = \langle \mathbf{a}, \mathbf{n}'_e \rangle \mathbf{n}'_e + \langle \mathbf{a}, \mathbf{n}'_y \rangle \mathbf{n}'_y.$$
(27)

Using the definition of \mathbf{n}'_e and the orthogonality between \mathbf{n}'_u and \mathbf{b}_{ex} , we have

$$\langle \operatorname{proj}(\mathbf{a}), \mathbf{b}_{ex} \rangle^2 = \langle \mathbf{a}, \mathbf{n}'_e \rangle^2 \|\mathbf{b}_{ex}\|_2^2 \stackrel{(a)}{=} \langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \stackrel{(8)}{\leq} \varepsilon,$$
 (28)

where (a) follows from the definition of \mathbf{n}'_e .

C.3. Characterizing the optimal a^{*} in (8)

Fix any $\varepsilon > 0$. For d = 2, it is convenient to work with polar coordinates. Let $\mathbf{b}_{ex} = r_e(\cos \theta_e, \sin \theta_e)$ for some $r_e > 0$ and $\theta_e \in [0, 2\pi]$. Similarly, let $\mathbf{b}_{yx} = r_y(\cos \theta_y, \sin \theta_y)$ for some $r_y > 0$ and $\theta_y \in [0, 2\pi]$. Let \mathbf{a}^* denote the set of optimal \mathbf{a} in (8), i.e.,

$$\mathbf{a}^{\star} = \arg \max_{\mathbf{a} \in \mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \text{ s.t } \left\langle \mathbf{a}, \mathbf{b}_{ex} \right\rangle^2 \le \varepsilon.$$
(29)

Then, the following lemma characterizes \mathbf{a}^* depending on the values of r_e, θ_e, θ_y , and ε .

Lemma 1. Let $\alpha \in [0, 2\pi]$ be such that $\cos \alpha = \sqrt{\varepsilon}/r_e$ and $\sin \alpha = \sqrt{1 - \varepsilon/r_e^2}$. Define the function $\mathbf{a} : [0, 2\pi] \to [-1, 1]^2$ such that $\mathbf{a}(\theta) = (\cos \theta, \sin \theta)$. Then,

Case 1.1.
$$\mathbf{a}^* = \{\mathbf{a}(\alpha + \theta_e), \mathbf{a}(\pi + \alpha + \theta_e)\} \iff \theta_y \in [\theta_e, \alpha + \theta_e] \cup [\pi + \theta_e, \pi + \alpha + \theta_e],$$

$$Case \ 1.2. \ \mathbf{a}^{\star} = \left\{ \mathbf{a}(\theta_y), \mathbf{a}(\pi + \theta_y) \right\} \iff \theta_y \in [\alpha + \theta_e, \pi - \alpha + \theta_e] \cup [\pi + \alpha + \theta_e, 2\pi - \alpha + \theta_e], and$$

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677 Figure 6. Visualizing QCQP in (8). Each point is shown in polar coordinates, i.e., a point (r, θ) denotes $(r \cos \theta, r \sin \theta)$. The point \mathbf{b}_{ex} 678 is shown in orange. The region enclosed by the solid red lines and red arcs is the feasible space in (8). The feasible space is shaded in 679 blue, red, and olive to represent Case 1.1., Case 1.2., and Case 1.3., respectively. The point \mathbf{b}_{yx} is shown in magenta and falls under Case 680 1.3. The optimal solution set \mathbf{a}^* consists of points $(1, \pi - \alpha + \theta_e)$ and $(1, -\alpha, +\theta_e)$.

$$Case \ 1.3. \ \mathbf{a}^{\star} = \left\{ \mathbf{a}(\pi - \alpha + \theta_e), \mathbf{a}(-\alpha + \theta_e) \right\} \iff \theta_y \in \left[\pi - \alpha + \theta_e, \pi + \theta_e\right] \cup \left[-\alpha + \theta_e, \theta_e\right].$$

687 *Proof.* Consider any $\mathbf{b}_{ex} = r_e(\cos\theta_e, \sin\theta_e)$ (shown in orange in Figure 6) and $\varepsilon > 0$. Let $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex}) \triangleq {\mathbf{a} : \langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \le 1}$ 688 ε and $\mathbf{a} \in \mathcal{B}(0,1)$ be the feasible space in (8), i.e., the set of all \mathbf{a} satisfying the fairness constraint in (8) w.r.t \mathbf{b}_{ex} . Then, 689 $\Lambda_{a}(\mathbf{b}_{ex})$ is the region enclosed by the solid red lines and red arcs in Figure 6, i.e., the union of the regions shaded in various 690 colors. To obtain this region, consider any $\mathbf{a} \in \mathcal{B}(0,1)$. Then, ensuring $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$ is equivalent to ensuring that the 691 projection of **a** on the line joining the origin and \mathbf{b}_{ex} (shown in dotted orange in Figure 6) is no more than $\sqrt{\varepsilon}/r_e$. Therefore, 692 we drop perpendiculars to this line at a distance $\sqrt{\varepsilon}/r_e$ on either side of the origin. We obtain the intersection of these 693 perpendiculars and $\mathcal{B}(0,1)$ by standard algebra and trigonometry (these points are shown in Figure 6 in polar coordinates), 694 and the enclosed region is $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex})$. 695

Now, consider any $\mathbf{b}_{yx} = r_y(\cos\theta_y, \sin\theta_y)$ (shown in magenta in Figure 6). Then, obtaining \mathbf{a}^* is equivalent to obtaining every $\mathbf{a} \in \Lambda_{\mathbf{a}}(\mathbf{b}_{ex})$ such that the projection of \mathbf{a} on the line joining the origin and \mathbf{b}_{yx} (shown in dotted magenta in Figure 6) is maximized. As a result, every $\mathbf{a} \in \mathbf{a}^*$ lies on the boundary of this region. To obtain \mathbf{a}^* , we drop perpendiculars to the line joining the origin and \mathbf{b}_{yx} from every \mathbf{a} on the boundary of this region, e.g., we show two such perpendiculars in solid blue in Figure 6. We add \mathbf{a} to \mathbf{a}^* if the distance between the origin and the point where the corresponding perpendicular intersects the line is maximum. It is straightforward to verify that Case 1.1., Case 1.2., and Case 1.3. correspond to the regions shaded in blue, red, and olive, respectively (see Figure 6).

D. Proof of Proposition 2: d = 2 suffices for robust QCQP

It suffices to show that the projection of any $\mathbf{a} \in \mathcal{B}(0, 1)$, satisfying the constraint $\langle \mathbf{a}, \mathbf{b} \rangle^2 \leq \varepsilon$ for all $\mathbf{b} \in \mathcal{B}(\hat{\mathbf{b}}_{ex}, \tau)$, onto the subspace spanned by \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$ preserves the value of the objective and continues to satisfy the constraint. Fix some **a** in the feasible space. Let $\operatorname{proj}(\mathbf{a})$ denote the projection of **a** on the subspace spanned by \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$. The proof that the value of the objective remains preserved is analogous to the proof of Proposition 1 (see Appendix C.2).

The constraint is satisfied. It suffices to show $\operatorname{proj}(\mathbf{a}) \in \mathcal{B}(0,1)$ and $\langle \operatorname{proj}(\mathbf{a}), \mathbf{b} \rangle^2 \leq \varepsilon$ for all $\mathbf{b} \in \mathcal{B}(\widehat{\mathbf{b}}_{ex}, \tau)$. It is easy to see $\operatorname{proj}(\mathbf{a}) \in \mathcal{B}(0,1)$ because $\|\operatorname{proj}(\mathbf{a})\|_2 \leq \|\mathbf{a}\|_2$. Now, consider a set of orthonormal basis $\{\mathbf{n}_e, \mathbf{n}_y\} \triangleq \{\frac{\widehat{\mathbf{b}}_{ex}}{\|\widehat{\mathbf{b}}_{ex}\|_2}, \mathbf{n}_y\}$

spanned by $\hat{\mathbf{b}}_{ex}$ and \mathbf{b}_{yx} where $\mathbf{n}_e = \frac{\hat{\mathbf{b}}_{ex}}{\|\hat{\mathbf{b}}_{ex}\|_2}$ and \mathbf{n}_y is chosen to be orthogonal to \mathbf{n}_e . Then,

$$\operatorname{proj}(\mathbf{a}) = \langle \mathbf{a}, \mathbf{n}_e \rangle \mathbf{n}_e + \langle \mathbf{a}, \mathbf{n}_y \rangle \mathbf{n}_y.$$
(30)

Using the definition of \mathbf{n}_e and the orthogonality between \mathbf{n}_y and $\hat{\mathbf{b}}_{ex}$, we have

$$\langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \rangle = \langle \mathbf{a}, \mathbf{n}_e \rangle \| \widehat{\mathbf{b}}_{ex} \|_2 \stackrel{(a)}{=} \langle \mathbf{a}, \widehat{\mathbf{b}}_{ex} \rangle,$$
 (31)

where (a) follows from the definition of \mathbf{n}_e .

Now, for every $\mathbf{b} \in \mathcal{B}(\widehat{\mathbf{b}}_{ex}, \tau)$, we have $\mathbf{b} = \widehat{\mathbf{b}}_{ex} + \Delta \mathbf{b}$, where $\|\Delta \mathbf{b}\| \leq \tau$. Then,

$$\left\langle \mathbf{a}, \mathbf{b} \right\rangle = \left\langle \mathbf{a}, \widehat{\mathbf{b}}_{ex} + \Delta \mathbf{b} \right\rangle = \left\langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \right\rangle + \left\langle \mathbf{a}, \Delta \mathbf{b} \right\rangle \stackrel{(31)}{=} \left\langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \right\rangle + \left\langle \mathbf{a}, \Delta \mathbf{b} \right\rangle \leq \left\langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \right\rangle + \|\mathbf{a}\|_{2} \tau \quad (32)$$

$$\leq \sqrt{\varepsilon},$$
 (33)

where (a) follows because $\langle \mathbf{a}, \mathbf{b} \rangle^2 \leq \varepsilon$. Similarly, we can show $\langle \mathbf{a}, \mathbf{b} \rangle \geq -\sqrt{\varepsilon}$. Now, we can bound $\langle \operatorname{proj}(\mathbf{a}), \mathbf{b} \rangle$ as follows:

$$\langle \operatorname{proj}(\mathbf{a}), \mathbf{b} \rangle = \langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} + \Delta \mathbf{b} \rangle = \langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \rangle + \langle \operatorname{proj}(\mathbf{a}), \Delta \mathbf{b} \rangle \leq \langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \rangle + \| \operatorname{proj}(\mathbf{a}) \|_2 \tau$$
(34)

$$\stackrel{(a)}{\leq} \langle \operatorname{proj}(\mathbf{a}), \widehat{\mathbf{b}}_{ex} \rangle + \|\mathbf{a}\|_2 r \stackrel{(33)}{\leq} \sqrt{\varepsilon}, \quad (35)$$

where (a) follows because $\|\text{proj}(\mathbf{a})\|_2 \leq \|\mathbf{a}\|_2$. Similarly, we can show $\langle \text{proj}(\mathbf{a}), \mathbf{b} \rangle \geq -\sqrt{\varepsilon}$ completing the proof.

E. Proof of Theorem B.2: Robust QCQP with infinite constraints

Fix any $\hat{r}_e \ge 0$, $\hat{\theta}_e \in [0, 2\pi]$, $\Delta \ge 0$, and $\phi \in [0, \pi/2]$. For any **b**, let $\Lambda_{\mathbf{a}}(\mathbf{b}) \triangleq \{\mathbf{a} : \langle \mathbf{a}, \mathbf{b} \rangle^2 \le \varepsilon$ and $\mathbf{a} \in \mathcal{B}(0, 1)\}$ be the set of all **a** satisfying the fairness constraint in (8) w.r.t **b**. For a given **b**, $\Lambda_{\mathbf{a}}(\mathbf{b})$ can be constructed as described in the proof of Lemma 1 in Appendix C.3. See Figure 6 for reference.

Now, to show that the optimization problem in (11) is equivalent to the optimization problem in (12), it suffices to show the corresponding constraints are equivalent, i.e.,

$$\bigcap_{\mathbf{b}\in\mathcal{A}(\Delta,\phi)}\Lambda_{\mathbf{a}}(\mathbf{b}) = \bigcap_{\mathbf{b}\in\overline{\mathcal{A}}(\Delta,\phi)}\Lambda_{\mathbf{a}}(\mathbf{b}).$$
(36)

Consider any $\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{A}(\Delta, \phi)$ of the form: $\mathbf{b}_1 = r_1(\cos \theta, \sin \theta)$ and $\mathbf{b}_2 = r_2(\cos \theta, \sin \theta)$ for some r_1, r_2, θ such that $|r_1 - \hat{r}_e| \leq \Delta$, $|r_2 - \hat{r}_e| \leq \Delta$, and $|\theta - \hat{\theta}_e| \leq \phi$. Without loss of generality, let $r_1 \geq r_2$. Now, it is straightforward to see that $\Lambda_{\mathbf{a}}(\mathbf{b}_1) \subseteq \Lambda_{\mathbf{a}}(\mathbf{b}_2)$. Therefore, (36) follows from the definition of $\overline{\mathcal{A}}(\Delta, \phi)$.

E.1. Characterizing optimal a in (11)

Fix any $\varepsilon > 0$. Let $\widehat{\mathbf{b}}_{ex} = \widehat{r}_e(\cos \widehat{\theta}_e, \sin \widehat{\theta}_e)$ for some $\widehat{r}_e > 0$ and $\widehat{\theta}_e \in [0, 2\pi]$. Consider some $\Delta \ge 0$ and $\phi \in [0, \pi/2]$ such that $\mathbf{b}_{ex} \in \mathcal{A}(\Delta, \phi)$ with probability at least $1 - \delta$ where $\mathcal{A}(\Delta, \phi)$ is as defined in (10). Let $\mathbf{b}_{yx} = r_y(\cos \theta_y, \sin \theta_y)$ for some $r_y > 0$ and $\theta_y \in [0, 2\pi]$. Let $\overline{\mathbf{a}}$ denote the set of optimal \mathbf{a} in (11), i.e.,

$$\overline{\mathbf{a}} = \arg \max_{\mathbf{a} \in \mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \text{ s.t } \left\langle \mathbf{a}, \mathbf{b} \right\rangle^2 \le \varepsilon \text{ for all } \mathbf{b} \in \mathcal{A}(\Delta, \phi).$$
(37)

Then, the following Lemma characterizes $\overline{\mathbf{a}}$ depending on the values of $\widehat{\mathbf{b}}_{ex}$, \mathbf{b}_{yx} , Δ , ϕ , and ε .

Lemma 2. Let $\overline{\alpha} \in [0, 2\pi]$ be such that $\cos \overline{\alpha} = \sqrt{\varepsilon}/(\widehat{r}_e + \Delta)$ and $\sin \overline{\alpha} = \sqrt{1 - \varepsilon/(\widehat{r}_e + \Delta)^2}$. Define the function $\mathbf{a} : [0, 2\pi] \to [-1, 1]^2$ such that $\mathbf{a}(\theta) = (\cos \theta, \sin \theta)$. Then,

1. If
$$\sqrt{\varepsilon} \ge (\widehat{r}_e + \Delta) \sin \phi$$
:



Figure 7. Visualizing QCQP in (11) for $\varepsilon = 0.9$, $\hat{r}_e = 1.6$, and $\hat{\theta}_e = 0$. We set $\Delta = 0.2$ and $\phi = 15$ for panel (a), and $\Delta = 0.4$ and $\phi = 30$ for panel (b). Each point is shown in polar coordinates, i.e., a point (r, θ) denotes $(r \cos \theta, r \sin \theta)$. The annular sector $\mathcal{A}(\Delta, \phi)$ is shown in orange. The region enclosed by the solid blue lines and blue arcs is the feasible space in (11). In Panel (a), the feasible space is shaded in green, blue, red, and olive to represent Case 2.1., Case 2.2., Case 2.3., and Case 2.4. respectively. In Panel (b), the feasible space is shaded in green and blue to represent Case 3.1. and Case 3.2., respectively.

Proof. For any **b**, let $\Lambda_{\mathbf{a}}(\mathbf{b}) \triangleq \{\mathbf{a} : \langle \mathbf{a}, \mathbf{b} \rangle^2 \leq \varepsilon \text{ and } \mathbf{a} \in \mathcal{B}(0,1)\}$ be the set of all **a** satisfying the fairness constraint in (8) w.r.t **b**. Then, from (36), the constraint in (37) is equivalent to $\bigcap_{\mathbf{b} \in \overline{\mathcal{A}}(\Delta,\phi)} \Lambda_{\mathbf{a}}(\mathbf{b})$. Consider any $\mathbf{b}_1 \in \overline{\mathcal{A}}(\Delta,\phi)$. Then, $\mathbf{b}_1 = (\hat{r}_e + \Delta)(\cos\theta, \sin\theta)$ for some $|\theta - \hat{\theta}_e| \leq \phi$. Further, $\Lambda_{\mathbf{a}}(\mathbf{b}_1)$ can be constructed as described in the proof of Lemma 1 in Appendix C.3. See Figure 6 for reference. From standard algebra and trigonometry, it is easy to see the two straight lines forming the boundary of $\Lambda_{\mathbf{a}}(\mathbf{b}_1)$ in Figure 6 intersect the y-axis at points $(0, \frac{\sqrt{\varepsilon}}{(\hat{r}_e + \Delta)\sin\theta})$ and $(0, \frac{-\sqrt{\varepsilon}}{(\hat{r}_e + \Delta)\sin\theta})$. Then, depending on whether these points lie inside or outside $\mathcal{B}(0, 1)$ when $\theta = \phi$, we have two cases.

1. Suppose $\sqrt{\varepsilon} \ge (\hat{r}_e + \Delta) \sin \phi$. Then, it is straightforward to see that $\bigcap_{\mathbf{b} \in \overline{\mathcal{A}}(\Delta,\phi)} \Lambda_{\mathbf{a}}(\mathbf{b})$ is the region enclosed by the solid blue lines and blue arcs in Figure 7(a) (see Figure 4(a) for reference), i.e., the union of the regions shaded in various colors. The rest of the proof is similar to the proof of Lemma 1. Lastly, it is straightforward to verify that Case 2.1., Case 2.2., Case 2.3., and Case 2.4. correspond to the regions shaded in green, blue, red, and olive, respectively.

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2. Suppose √ε ≤ (r̂_e + Δ) sin φ. Then, it is straightforward to see that ∩_{b∈A(Δ,φ)}Λ_a(b) is the region enclosed by the solid blue lines and blue arcs in Figure 7(b) (see Figure 4(b) for reference), i.e., the union of the regions shaded in various colors. The rest of the proof is similar to the proof of Lemma 1. Lastly, it is straightforward to verify that Case 3.1. and Case 3.2. correspond to the regions shaded in green and blue, respectively.

F. Proof of Theorem B.3: Robust QCQP with 3 constraints

Fix any $\hat{r}_e \ge 0$, $\theta_e \in [0, 2\pi]$, $\Delta \ge 0$, and $\phi \in [0, \pi/2]$. For any **b**, let $\Lambda_{\mathbf{a}}(\mathbf{b}) \triangleq \{\mathbf{a} : \langle \mathbf{a}, \mathbf{b} \rangle^2 \le \varepsilon$ and $\mathbf{a} \in \mathcal{B}(0, 1)\}$ be the set of all **a** satisfying the fairness constraint in (8) w.r.t **b**. For a given **b**, $\Lambda_{\mathbf{a}}(\mathbf{b})$ can be constructed as described in the proof of Lemma 1 in Appendix C.3. See Figure 6 for reference.

First, we note that the straight lines in the feasible space $\bigcap_{\mathbf{b}\in\overline{\mathcal{A}}(\Delta,\phi)}\Lambda_{\mathbf{a}}(\mathbf{b})$ in Figure 7(a) and Figure 7(b) are generated by the extreme points of the arc $\overline{\mathcal{A}}(\Delta,\phi)$. These extreme points are precisely $\mathbf{b}_{ex}^{(2)}$ and $\mathbf{b}_{ex}^{(3)}$. Therefore, the same straight lines also arise in the feasible space $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(1)}) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(3)}) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(3)})$ in Figure 8(a) and Figure 8(b). Second, we note that the arcs in the feasible space $\bigcap_{\mathbf{b}\in\overline{\mathcal{A}}(\Delta,\phi)}\Lambda_{\mathbf{a}}(\mathbf{b})$ in Figure 7(a) and Figure 7(b) are at a distance $\sqrt{\varepsilon}/(\hat{r}_e + \Delta)$ away from the origin. The point $\mathbf{b}_{ex}^{(1)}$ is chosen precisely such that the space $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(1)})$ intersects each of these arcs at its extreme points (see Figure 8(a) and Figure 8(b)). Therefore, it is easy to see that the boundary of the feasible space $\bigcap_{\mathbf{b}\in\overline{\mathcal{A}}(\Delta,\phi)}\Lambda_{\mathbf{a}}(\mathbf{b})$ is not closer to the origin than the boundary of the feasible space $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(1)}) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(2)}) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(3)})$. This completes the proof.



(a) Feasible space for $\sqrt{\varepsilon} > (\hat{r}_e + \Delta) \sin \phi$. (b) Feasible space for $\sqrt{\varepsilon} < (\hat{r}_e + \Delta) \sin \phi$.

Figure 8. Visualizing QCQP in (13) for $\varepsilon = 0.9$, $\hat{r}_e = 1.6$, and $\hat{\theta}_e = 0$. We set $\Delta = 0.2$ and $\phi = 15$ for panel (a), and $\Delta = 0.4$ and $\phi = 30$ for panel (b). Each point is shown in polar coordinates, i.e., a point (r, θ) denotes $(r \cos \theta, r \sin \theta)$. The annular sector $\mathcal{A}(\Delta, \phi)$ is shown in orange. The shaded region is the feasible space in (13). The points $\mathbf{b}_{ex}^{(1)}$, $\mathbf{b}_{ex}^{(2)}$, and $\mathbf{b}_{ex}^{(3)}$ are shown in magenta, blue and green, respectively.

F.1. Characterizing optimal a in (13)

Fix any $\varepsilon > 0$. Let $\hat{\mathbf{b}}_{ex} = \hat{r}_e(\cos \hat{\theta}_e, \sin \hat{\theta}_e)$ for some $\hat{r}_e > 0$ and $\hat{\theta}_e \in [0, 2\pi]$. Consider some $\Delta \ge 0$ and $\phi \in [0, \pi/2]$ such that $\mathbf{b}_{ex} \in \mathcal{A}(\Delta, \phi)$ where $\mathcal{A}(\Delta, \phi)$ is as defined in (10). Let $\mathbf{b}_{yx} = r_y(\cos \theta_y, \sin \theta_y)$ for some $r_y > 0$ and $\theta_y \in [0, 2\pi]$. Let $\overline{\mathbf{a}}$ denote the set of optimal \mathbf{a} in (13), i.e.,

$$\overline{\mathbf{a}} = \arg \max_{\mathbf{a} \in \mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \text{ s.t } \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(1)} \right\rangle^2 \le \varepsilon, \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(2)} \right\rangle^2 \le \varepsilon, \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(3)} \right\rangle^2 \le \varepsilon.$$
(38)

Then, the following Lemma characterizes $\overline{\mathbf{a}}$ depending on the values of $\widehat{\mathbf{b}}_{ex}$, \mathbf{b}_{yx} , Δ , ϕ , and ε . The proof is similar to the proof of Lemma 2 and is omitted for brevity.

Lemma 3. Let $\overline{\alpha} \in [0, 2\pi]$ be such that $\cos \overline{\alpha} = \sqrt{\varepsilon}/(\widehat{r}_e + \Delta)$ and $\sin \overline{\alpha} = \sqrt{1 - \varepsilon/(\widehat{r}_e + \Delta)^2}$. Define the function $\mathbf{a} : [0, 2\pi] \to [-1, 1]^2$ such that $\mathbf{a}(\theta) = (\cos \theta, \sin \theta)$. Then,

$$\begin{aligned} &I. \ If \sqrt{\varepsilon} \geq (\widehat{r}_e + \Delta) \sin \phi: \\ &Case \ 4.1. \ \ \overline{\mathbf{a}} = \frac{\sqrt{\varepsilon}}{\widehat{r}_e + \Delta} \big\{ \mathbf{a}(\widehat{\theta}_e + \phi), \mathbf{a}(\pi + \widehat{\theta}_e + \phi) \big\} \iff \theta_y \in [\widehat{\theta}_e, \widehat{\theta}_e + \phi] \cup [\pi + \widehat{\theta}_e, \pi + \widehat{\theta}_e + \phi], \\ &Case \ 4.2. \ \ \overline{\mathbf{a}} = \big\{ \mathbf{a}(\widehat{\theta}_e + \phi + \overline{\alpha}), \mathbf{a}(\pi + \widehat{\theta}_e + \phi + \overline{\alpha}) \big\} \iff \theta_y \in [\widehat{\theta}_e + \phi, \widehat{\theta}_e + \phi + \overline{\alpha}] \cup [\pi + \widehat{\theta}_e + \phi, \pi + \widehat{\theta}_e + \phi + \overline{\alpha}], \\ &Case \ 4.3. \ \ \overline{\mathbf{a}} = \big\{ \mathbf{a}(\theta_y), \mathbf{a}(\pi + \theta_y) \big\} \iff \theta_y \in [\widehat{\theta}_e + \phi + \overline{\alpha}, \pi + \widehat{\theta}_e - \phi - \overline{\alpha}] \cup [\pi + \widehat{\theta}_e + \phi + \overline{\alpha}, 2\pi + \widehat{\theta}_e - \phi - \overline{\alpha}], \\ &Case \ 4.4. \ \ \overline{\mathbf{a}} = \big\{ \mathbf{a}(\pi + \widehat{\theta}_e - \phi - \overline{\alpha}), \mathbf{a}(\widehat{\theta}_e - \phi - \overline{\alpha}) \big\} \iff \theta_y \in [\pi + \widehat{\theta}_e - \phi - \overline{\alpha}, \pi + \widehat{\theta}_e - \phi] \cup [\widehat{\theta}_e - \phi - \overline{\alpha}, \widehat{\theta}_e - \phi], \\ &Case \ 4.5. \ \ \overline{\mathbf{a}} = \frac{\sqrt{\varepsilon}}{\widehat{r}_e + \Delta} \big\{ \mathbf{a}(\widehat{\theta}_e - \phi), \mathbf{a}(\pi + \widehat{\theta}_e - \phi) \big\} \iff \theta_y \in [\pi + \widehat{\theta}_e - \phi, \pi + \widehat{\theta}_e] \cup [\widehat{\theta}_e - \phi, \widehat{\theta}_e]. \end{aligned}$$

880 2. If $\sqrt{\varepsilon} \le (\hat{r}_e + \Delta) \sin \phi$:

$$\textit{Case 5.1.} \ \, \overline{\mathbf{a}} = \frac{\sqrt{\varepsilon}}{\widehat{r}_e + \Delta} \big\{ \mathbf{a}(\widehat{\theta}_e + \phi), \mathbf{a}(\pi + \widehat{\theta}_e + \phi) \big\} \iff \theta_y \in [\widehat{\theta}_e, \widehat{\theta}_e + \phi] \cup [\pi + \widehat{\theta}_e, \pi + \widehat{\theta}_e + \phi],$$

$$Case 5.2. \ \overline{\mathbf{a}} = \frac{\sqrt{\varepsilon}}{(\widehat{r}_e + \Delta)\sin\phi} \left\{ \mathbf{a}(\frac{\pi}{2} + \widehat{\theta}_e), \mathbf{a}(\frac{-\pi}{2} + \widehat{\theta}_e) \right\} \iff \theta_y \in [\widehat{\theta}_e + \phi, \pi + \widehat{\theta}_e - \phi] \cup [\pi + \widehat{\theta}_e + \phi, 2\pi + \widehat{\theta}_e - \phi],$$

$$Case 5.3. \ \overline{\mathbf{a}} = \frac{\sqrt{\varepsilon}}{\widehat{r}_e + \Delta} \left\{ \mathbf{a}(\widehat{\theta}_e - \phi), \mathbf{a}(\pi + \widehat{\theta}_e - \phi) \right\} \iff \theta_y \in [\pi + \widehat{\theta}_e - \phi, \pi + \widehat{\theta}_e] \cup [\widehat{\theta}_e - \phi, \widehat{\theta}_e].$$

G. Analyzing the power of labeled sensitive attributes

In this section, we consider the case where uncertainty only stems from sensitive attributes missing at random, and the uncertainty can be improved by collecting more labeled sensitive attributes. We characterize the difference in the optimal objectives in (8) and (13) to obtain the power of each new labeled sensitive attribute.

Fix any $\varepsilon > 0$. For any n, let $\hat{\mathbf{b}}_{ex}(n) \triangleq \hat{r}_e(n)(\cos \hat{\theta}_e(n), \sin \hat{\theta}_e(n))$ denote the estimate of \mathbf{b}_{ex} as a function of n and let $\Delta(n) \ge 0$ and $\phi(n) \in [0, \pi/2]$ denote the corresponding uncertainty parameters as a function of n, i.e., $\mathbf{b}_{ex} = r_e(\cos \theta_e, \sin \theta_e) \in \mathcal{A}(\Delta(n), \phi(n))$, with probability $1 - \delta$. Let $\Psi(n)$ denote the difference in the optimal objectives in (8) and (13), i.e., $\Psi(n) \triangleq \Psi - \Psi_2(n)$ where

$$\Psi \triangleq \max_{\mathbf{a} \in \mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \quad \text{s.t} \quad \left\langle \mathbf{a}, \mathbf{b}_{ex} \right\rangle^2 \le \varepsilon$$
(39)

and

$$\Psi_2(n) \triangleq \max_{\mathbf{a} \in \mathcal{B}(0,1)} \left\langle \mathbf{a}, \mathbf{b}_{yx} \right\rangle^2 \quad \text{s.t} \quad \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(1)}(n) \right\rangle^2 \le \varepsilon, \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(2)}(n) \right\rangle^2 \le \varepsilon, \left\langle \mathbf{a}, \mathbf{b}_{ex}^{(3)}(n) \right\rangle^2 \le \varepsilon.$$
(40)

Then, the following lemma characterizes $\Psi(n)$ as a function of n when $\theta_e \ge \hat{\theta}_e(n)$. The characterization when $\theta_e \le \hat{\theta}_e(n)$ can be obtained analogously. The proof follows from Lemma 1 and Lemma 3, and is omitted for brevity.

Lemma 4. Let $\alpha, \overline{\alpha}(n) \in [0, 2\pi]$ be such that (a) $\cos \alpha = \sqrt{\varepsilon}/r_e$ and $\sin \alpha = \sqrt{1 - \varepsilon/r_e^2}$ and (b) $\cos \overline{\alpha}(n) = \sqrt{\varepsilon}/(\widehat{r}_e(n) + \Delta(n))$ and $\sin \overline{\alpha}(n) = \sqrt{1 - \varepsilon/(\widehat{r}_e(n) + \Delta(n))^2}$. Then,

1. If
$$\sqrt{\varepsilon} \ge (\widehat{r}_e(n) + \Delta(n)) \sin \phi(n)$$
:

Case A.
$$\Psi(n) = \cos(\alpha + \theta_e - \theta_y) - \cos\overline{\alpha}(n)\cos(\widehat{\theta}_e(n) + \phi(n) - \theta_y) \text{ when } \theta_y \in [\theta_e, \widehat{\theta}_e(n) + \phi(n)] \cup [\pi + \theta_e, \pi + \widehat{\theta}_e(n) + \phi(n)],$$

$$\begin{aligned} \text{Case B.} \quad \Psi(n) = &\cos\left(\alpha + \theta_e - \theta_y\right) - \cos\left(\widehat{\theta}_e(n) + \overline{\alpha}(n) + \phi(n) - \theta_y\right) \text{ when } \theta_y \in [\widehat{\theta}_e(n) + \phi(n), \alpha + \theta_e] \cup [\pi + \widehat{\theta}_e(n) + \phi(n), \pi + \alpha + \theta_e], \end{aligned}$$

$$\begin{aligned} \text{Case C.} \quad \Psi(n) = 1 - \cos\left(\widehat{\theta}_e(n) + \overline{\alpha}(n) + \phi(n) - \theta_y\right) \text{ when } \theta_y \in [\alpha + \theta_e, \widehat{\theta}_e(n) + \phi(n) + \overline{\alpha}(n)] \cup [\pi + \alpha + \theta_e, \pi + \widehat{\theta}_e(n) + \phi(n) + \overline{\alpha}(n)], \end{aligned}$$

$$\begin{array}{ll} \begin{array}{l} 919\\ 920\\ 921\\ 921\\ 921\\ \end{array} \quad \begin{array}{l} Case \ D. \quad \Psi(n) = 0 \ when \ \theta_y \in [\widehat{\theta}_e(n) + \phi(n) + \overline{\alpha}(n), \pi + \widehat{\theta}_e(n) - \phi(n) - \overline{\alpha}(n)] \cup [\pi + \widehat{\theta}_e(n) + \phi(n) + \overline{\alpha}(n), 2\pi + \widehat{\theta}_e(n) - \phi(n) - \overline{\alpha}(n)], \end{array}$$

$$\begin{array}{ll} \textit{Case E.} \quad \Psi(n) = 1 - \cos\left(\theta_y + \overline{\alpha}(n) + \phi(n) - \widehat{\theta}_e(n)\right) \textit{ when } \theta_y \in [\pi + \widehat{\theta}_e(n) - \phi(n) - \overline{\alpha}(n), \pi - \alpha + \theta_e] \cup [\widehat{\theta}_e(n) - \phi(n) - \overline{\alpha}(n), -\alpha + \theta_e], \end{array}$$

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$$Case \ F. \quad \Psi(n) = \cos(\theta_y + \alpha - \theta_e) - \cos(\theta_y + \overline{\alpha}(n) + \phi(n) - \widehat{\theta}_e(n)) \text{when } \theta_y \in [\pi - \alpha + \theta_e, \pi + \widehat{\theta}_e(n) - \phi(n)] \cup [-\alpha + \theta_e, \widehat{\theta}_e(n) - \phi(n)],$$

$$\begin{array}{l} \begin{array}{l} \begin{array}{l} 226\\ 927\\ 928 \end{array} \end{array} \quad \mathcal{C}ase \ \mathcal{G}. \quad \Psi(n) = \cos\left(\theta_y + \alpha - \theta_e\right) - \cos\overline{\alpha}(n)\cos\left(\theta_y + \phi(n) - \widehat{\theta}_e(n)\right) \text{ when } \theta_y \in \left[\pi + \widehat{\theta}_e(n) - \phi(n), \pi + \widehat{\theta}_e(n)\right] \cup \\ \left[\widehat{\theta}_e(n) - \phi(n), \widehat{\theta}_e(n)\right], \end{array}$$

$$\textit{Case H.} \quad \Psi(n) = \cos\left(\theta_y + \alpha - \theta_e\right) - \cos\overline{\alpha}(n)\cos\left(\theta_y - \phi(n) - \widehat{\theta}_e(n)\right) \textit{ when } \theta_y \in [\pi + \widehat{\theta}_e(n), \pi + \theta_e] \cup [\widehat{\theta}_e(n), \theta_e].$$

2. If
$$\sqrt{\varepsilon}s \leq (\hat{r}_e(n) + \Delta(n)) \sin \phi(n)$$
.

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935 936	Case B.	$\begin{split} \Psi(n) = &\cos\left(\alpha + \theta_e - \theta_y\right) - \cos\overline{\alpha}(n) \frac{\sin\left(\theta_y - \widehat{\theta}_e(n)\right)}{\sin\phi(n)} \text{ when } \theta_y \in [\widehat{\theta}_e(n) + \phi(n), \alpha + \theta_e] \cup [\pi + \widehat{\theta}_e(n) + \phi(n), \pi + \alpha + \theta_e], \end{split}$
937 938	Case C.	$\Psi(n) = 1 - \cos \overline{\alpha}(n) \frac{\sin \left(\theta_y - \widehat{\theta}_e(n)\right)}{\sin \phi(n)} \text{ when } \theta_y \in [\alpha + \theta_e, \pi - \alpha + \theta_e] \cup [\pi + \alpha + \theta_e, 2\pi - \alpha + \theta_e],$
939 940	Case D.	$\Psi(n) = \cos\left(\theta_y + \alpha - \theta_e\right) - \cos\overline{\alpha}(n) \frac{\sin\left(\theta_y - \widehat{\theta}_e(n)\right)}{\sin\phi(n)} \text{ when } \theta_y \in [\pi - \alpha + \theta_e, \pi + \widehat{\theta}_e(n) - \phi(n)] \cup [-\alpha + \theta_e, \widehat{\theta}_e(n) - \phi(n)] \cup $
941 942 943	Case E.	$\Psi(n) = \cos\left(\theta_y + \alpha - \theta_e\right) - \cos\overline{\alpha}(n)\cos\left(\theta_y + \phi(n) - \widehat{\theta}_e(n)\right) \text{ when } \theta_y \in \left[\pi + \widehat{\theta}_e(n) - \phi(n), \pi + \widehat{\theta}_e(n)\right] \cup \left[\widehat{\theta}_e(n) - \phi(n), \widehat{\theta}_e(n)\right].$
944	Case F.	$\Psi(n) = \cos\left(\theta_y + \alpha - \theta_e\right) - \cos\overline{\alpha}(n)\cos\left(\theta_y - \phi(n) - \widehat{\theta}_e(n)\right) \text{ when } \theta_y \in [\pi + \widehat{\theta}_e(n), \pi + \theta_e] \cup [\widehat{\theta}_e(n), \theta_e].$
945 946	Remark 1.	The cases above can be classified into one of the following three categories:
947 948 949 950 951 952	Category	<i>p</i> 1. Any uncertainty hurts: Here, the optimal performance of the robust QCQP in (11) matches the optimal performance in (8) only when all the uncertainty is removed, i.e., when $n \to N$. When the uncertainty parameters are not too large, i.e., $(\hat{r}_e(n) + \Delta(n)) \sin \phi(n) \le \sqrt{\varepsilon}$, Case A., Case B., Case F., Case G., and Case H. fall into this category. When the uncertainty parameters are large, i.e., $(\hat{r}_e(n) + \Delta(n)) \sin \phi(n) \ge \sqrt{\varepsilon}$, Case A., Case B., Case B., Case D., Case E., and Case F. fall into this category.
 953 954 955 956 957 958 	Category	2. Some uncertainty does not hurt: Here, the optimal performance of the robust QCQP in (11) matches the optimal performance in (8) when some uncertainty is removed, i.e., by collecting some additional labeled sensitive attributes. When the uncertainty parameters are not too large, i.e., $(\hat{r}_e(n) + \Delta(n)) \sin \phi(n) \le \sqrt{\epsilon}$, Case C. and Case E. fall into this category. When the uncertainty parameters are large, i.e., $(\hat{r}_e(n) + \Delta(n)) \sin \phi(n) \ge \sqrt{\epsilon}$, Case C. fall into this category.
 958 959 960 961 962 963 964 965 966 	Category	<i>b</i> 3. Uncertainty does not hurt: Here, the optimal performance of the robust QCQP in (11) matches the optimal performance in (8) without removing any uncertainty, i.e., without collecting any additional labeled sensitive attributes. This only happens in Case D. where $\Psi(n) = 0$ when the uncertainty parameters are not too large, i.e., $(\hat{r}_e(n) + \Delta(n)) \sin \phi(n) \le \sqrt{\varepsilon}$. Such a situation arises when \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$ are very close to being perpendicular to each other. In other words, when \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$ are close to being independent, our proposed robust QCQP achieves optimal performance while ensuring a strict fairness guarantee. We call such a phenomenon as "free fairness" (see Corollary 1).
967 968	H. Mono	tonic performance with new labeled sensitive attributes

As in the previous section, we consider the case where uncertainty only stems from sensitive attributes missing at random. Then, we show that the optimal objective in (13) either monotonically increases or coincides with the optimal objective in (8) whenever the uncertainty set monotonically decreases with n.

To that end, below, we express the uncertainty set $\mathcal{A}(\Delta, \phi)$ and the associated uncertainty parameters Δ and ϕ as a function of *n*.

Proposition 3 (Uncertainty as a function of *n*). Fix any $\delta > 0$. Given *n* samples $\{(e^{(i)}, \mathbf{x}^{(i)})\}_{i \in [n]}$ of (e, \mathbf{x}) , let

$$\widehat{\mathbf{b}}_{ex} \triangleq \mathbf{\Sigma}_{ee}^{-1/2} \widehat{\mathbf{\Sigma}}_{ex} \mathbf{\Sigma}_{xx}^{-1/2} \text{ with } \widehat{\mathbf{\Sigma}}_{ex} \triangleq \frac{1}{n} \sum_{i \in [n]} e^{(i)} \boldsymbol{x}^{(i)}.$$
(41)

980 Then, $\mathbf{b}_{ex} \in \mathcal{B}(\widehat{\mathbf{b}}_{ex}, \tau(n))$, with probability $1 - \delta$, where 981

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$$\tau(n) \triangleq \|\boldsymbol{\Sigma}_{xx}^{-1/2}\|_2 \frac{c\sqrt{\sigma_e} \max_{i \in [d]} \{\sigma_i\}}{n\sqrt{d}} \log \frac{4}{\delta},\tag{42}$$

984 985 with $\sigma_1^2, \sigma_2^2, \sigma_e^2$ denoting variances of $\mathbf{x}_1, \mathbf{x}_2$, e, respectively, and c is a universal constant. Further, for d = 2, $\mathbf{b}_{ex} \in \mathcal{A}(\Delta(n), \phi(n))$, with probability at least $1 - \delta$, where

$$\phi(n) \triangleq \sin^{-1}\left(\frac{\tau(n)}{\|\widehat{\mathbf{b}}_{ex}\|_2}\right) and \ \Delta(n) \triangleq \tau(n).$$
(43)

Proof. For any sub-Gaussian (sub-exponential) random variable, we denote its sub-Gaussian (sub-exponential) norm by $\|\cdot\|_{\Psi_2} (\|\cdot\|_{\Psi_1})$. For every $i \in [d]$, we note that x_i is a sub-Gaussian random variable with sub-Gaussian norm $\|x_i\|_{\Psi_2} = c\sigma_i$ where $\sigma_i^2 = \Sigma_{x_i x_i}$. Similarly, e is a sub-Gaussian random variable with sub-Gaussian norm $\|e\|_{\Psi_2} = c\sigma_e$ where $\sigma_e^2 = \Sigma_{ee}$. Then, from Vershynin (2018, Lemma. 2.7.7), for every $i \in [d]$, $x_i e$ is a sub-exponential random variable with sub-exponential norm $\|x_i e\|_{\Psi_1} = c\sigma_i\sigma_e$. Therefore, from Bernstein's inequality (Vershynin, 2018, Corollary 2.8.3), with probability $1 - \delta$, we have $\|\widehat{\Sigma}_{ex} - \Sigma_{ex}\|_{\infty} \leq t$ whenever $t \geq \frac{c\sigma_e \max_{i \in [d]} \sigma_i}{n} \log \frac{2d}{\delta}$. As a result, we have $\|\widehat{\Sigma}_{ex} - \Sigma_{ex}\|_2 \leq t$ with probability $1 - \delta$ whenever $t \geq \frac{c\sigma_e \max_{i \in [d]} \sigma_i}{n\sqrt{d}} \log \frac{2d}{\delta}$. Conditioning on this event, we have, with probability $1 - \delta$,

$$\|\widehat{\mathbf{b}}_{ex} - \mathbf{b}_{ex}\|_{2} \stackrel{(a)}{=} \|\mathbf{\Sigma}_{ee}^{-1/2} (\widehat{\mathbf{\Sigma}}_{ex} - \mathbf{\Sigma}_{ex}) \mathbf{\Sigma}_{xx}^{-1/2}\|_{2} \stackrel{(b)}{\leq} \frac{1}{\sqrt{\sigma_{e}}} \|\widehat{\mathbf{\Sigma}}_{ex} - \mathbf{\Sigma}_{ex}\|_{2} \|\mathbf{\Sigma}_{xx}^{-1/2}\|_{2} \le \frac{1}{\sqrt{\sigma_{e}}} \|\mathbf{\Sigma}_{xx}^{-1/2}\|_{2}$$
(44)

1001 where (a) follows from Definition 2 and (b) follows because induced matrix norms are sub-multiplicative. Then, letting 1002 d = 2 and $\tau(n) = \sum_{ee}^{-1/2} \|\sum_{xx}^{-1/2}\|_{\infty} t$, it is easy to verify $\mathbf{b}_{ex} \in \mathcal{A}(\Delta(n), \phi(n))$ whenever

$$\Delta(n) = \tau(n) \quad \text{and} \quad \phi(n) = \sin^{-1} \left(\frac{\tau(n)}{\|\widehat{\mathbf{b}}_{ex}\|_2} \right). \tag{45}$$

¹⁰⁰⁸ H.1. Monotonic performance with decrease in uncertainty set

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Now, in the following theorem, we show that the optimal objective of the robust QCQP with 3 constraints in (13) either monotonically increases or coincides with the optimal objective of the QCQP in (8) whenever the uncertainty set $\mathcal{A}(\Delta(n), \phi(n))$ monotonically decreases with *n*.

1013 **Theorem H.1** (Monotonic performance). For any number of labeled sensitive attributes n, let $\mathcal{A}(\Delta(n), \phi(n))$ denote the 1014 uncertainty set containing \mathbf{b}_{ex} . Let $\overline{\mathbf{a}}(n)$ denote the optimal solution \mathbf{a} in (13) as a function of n. If $\mathcal{A}(\Delta(n+1), \phi(n+1)) \subset 1015 \quad \mathcal{A}(\Delta(n), \phi(n))$, then

$$\left\langle \overline{\mathbf{a}}(n+1), \mathbf{b}_{yx} \right\rangle^2 > \left\langle \overline{\mathbf{a}}(n), \mathbf{b}_{yx} \right\rangle^2 \quad or$$
(46)

$$\left\langle \overline{\mathbf{a}}(n+1), \mathbf{b}_{yx} \right\rangle^2 = \left\langle \overline{\mathbf{a}}(n), \mathbf{b}_{yx} \right\rangle^2 = \left\langle \mathbf{a}^{\star}, \mathbf{b}_{yx} \right\rangle^2,$$
(47)

where \mathbf{a}^* is the optimal solution of the QCQP in (8).

Fix any $\varepsilon > 0$. For any **b**, let $\Lambda_{\mathbf{a}}(\mathbf{b}) \triangleq \{\mathbf{a} : \langle \mathbf{a}, \mathbf{b} \rangle^2 \le \varepsilon \text{ and } \mathbf{a} \in \mathcal{B}(0, 1)\}$ be the set of all **a** satisfying the fairness constraint in (8) w.r.t **b**. For a given **b**, $\Lambda_{\mathbf{a}}(\mathbf{b})$ can be constructed as described in the proof of Lemma 1 in Appendix C.3. See Figure 6 for reference.

Now, using $\mathcal{A}(\Delta(n+1), \phi(n+1)) \subset \mathcal{A}(\Delta(n), \phi(n))$, it is easy to see that $\hat{r}_e(n+1) < \hat{r}_e(n)$, $|\hat{\theta}_e(n+1) + \phi(n+1)| < |\hat{\theta}_e(n+1) - \phi(n+1)| < |\hat{\theta}_e(n) - \phi(n)|$, and $|\cos \phi(n+1)| < |\cos \phi(n)|$. Then, it follows that the feasible space increases with n, i.e., $\Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(1)}(n+1)) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(2)}(n+1)) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(3)}(n+1)) \supset \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(1)}(n)) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(2)}(n)) \cap \Lambda_{\mathbf{a}}(\mathbf{b}_{ex}^{(2)}(n))$. As a result, the optimal objective in (13) at n + 1 is either less than the optimal objective in (13) at n or equal to the optimal objective in (13) at n + 1 is either less than the optimal objective in (13) at n or equal to the optimal objective in (13) at n, then it also equals the optimal objective of the QCQP in (8). This follows directly from the expressions for $\Psi(n)$ in various cases of Lemma 4.

I. Additional experimental results

¹⁰³⁵ In this section, we provide additional empirical results as well as more implementation details.

¹⁰³⁷ I.1. Additional results for Gaussian data

1039 **Covariance matrices.** First, we provide the covariance matrices of (x, y, e) used in the experiments in Appendix B.3:

$$\Sigma_{2}^{\text{gen}} = \begin{bmatrix} 1 & 0.1 & 0.5 & 0.4 \\ 0.1 & 1 & 0.5 & 0.25 \\ \hline 0.5 & 0.5 & 1 & 0.75 \\ \hline 0.4 & 0.25 & 0.75 & 1 \end{bmatrix} \qquad \Sigma_{2}^{\text{fair}} = \begin{bmatrix} 1 & 0.1 & 0.5 & 0.05 \\ 0.1 & 1 & 0.05 & 0.25 \\ \hline 0.5 & 0.05 & 1 & 0.75 \\ \hline 0.05 & 0.25 & 0.75 & 1 \end{bmatrix}$$

1045 The choice of Σ_2^{fair} is deliberate to demonstrate the free-fairness behavior. As pointed out in Category 3 of Remark 1, the 1046 free-fairness behavior may happen when \mathbf{b}_{yx} and $\hat{\mathbf{b}}_{ex}$ are almost perpendicular. It is easy to check that this can be achieved 1047 by the following steps, for example: (1) Strongly correlate *e* with x_1 , and *y* with x_2 . (2) Set the covariance of **x** to be close 1048 to the identity matrix.

1049 **Results for** d = 3. Next, we provide results for d = 3 where we use the following covariance matrices:

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	1	0.1	0.5	0.5	0.4] [1	0.1	0.5	0.5	0.05
	0.1	1	0.5	0.5	0.25		0.1	1	0.5	0.05	0.25
$\Sigma_3^{\text{gen}} =$	0.5	0.5	1	0.2	0.2	$\Sigma_3^{\text{fair}} =$	0.5	0.5	1	0.2	0.2
Ť	0.5	0.5	0.2	1	0.75	_	0.5	0.05	0.2	1	0.75
	0.4	0.25	0.2	0.75	1	ļ	0.05	0.25	0.2	0.75	1

The reasoning behind the choice of Σ_3^{fair} is the same as above. As we see in Figure 9 (a) and Figure 9 (b), Bootstrap-S has significantly less fairness violations compared to Baseline with a similar MSE even for d = 3.



1080 Figure 9. The performance of Bootstrap-S with $S \in \{3, 9, 27\}$ and Baseline for d = 3 and various Σ^* . In the left column, we 1081 plot the fraction of violations of the fairness constraint $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$ vs. n; in the middle column, we plot average MSE vs. n; in the 1082 right column, we plot the histogram of the value of $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2$ over 1,000 trials for n = 250.

A different choice of covariance matrices. Next, we provide results with a different choice of covariance matrices:

	- 1	0.1	05			1	0.1	0.5	0.5	0.01	1
$\Sigma_2^{\mathrm{a}} =$		1	0.0	0.01 0.25 0.75		0.1	1	0.5	0.01	0.25	
	$\frac{0.1}{0.5}$	0.01	0.01		$\Sigma_3^a =$	0.5	0.5	1	0.2	0.2	
	$\frac{0.5}{0.01}$	0.01	1		_	0.5	0.01	0.2	1	0.75	
	0.01	0.20	0.75			0.01	0.25	0.2	0.75	1	

1092 As we see in Figure 10, the behavior is similar as in other scenarios.

Results for uncertainty due to noise. In Appendix B.3, we considered uncertainty arising due to limited number of sensitive attributes. Here, we consider uncertainty arising due to noise. Specifically, we generate synthetic data using the covariance matrix Σ_2^{fair} and conduct two sets of experiments as follows. In the first experiment, we add independent zero-mean Gaussian noise, i.e., $\mathcal{N}(0, \sigma^2)$, with varying noise level, i.e., $\sigma = 1, \dots, 10$, to the sensitive attributes. Such kind of noise is common when the sensitive attributes need to be privatized. In the second experiment, we add independent zero-mean Gaussian noise, i.e., $\mathcal{N}(0, \sigma^2)$, with varying noise level, i.e., $\sigma = 1, \dots, 10$, to the sensitive attributes only if

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Figure 10. The performance of robust QCQP in (13) (for d = 2), Bootstrap-S with $S \in \{3, 9, 27\}$, and Baseline for various d and Σ^* . In the left column, we plot the fraction of violations of the fairness constraint $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$ vs. n; in the middle column, we plot average MSE vs. n; in the right column, we plot the histogram of the value of $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2$ over 1000 trials for n = 250.

the sensitive attribute magnitude is larger than a certain threshold, i.e., $|e| \ge \Sigma_{ee}$. Such data-dependent noise simulates noisy responses when sensitive attributes are collected via a survey, e.g., minority groups pretend to be the majority due to fear of discrimination. In our experiment, samples with sensitive attribute magnitudes larger than the threshold have noisy responses. We provide the results for both experiments in Figure 11 where we observe that Bootstrap-S achieves significantly fewer fairness violations than Baseline while maintaining comparable MSE.

I.2. Additional results for real-world data

Dataset description. The Adult data (Lantz, 2019) (https://archive.ics.uci.edu/ml/datasets/adult) considers predicting whether an individual's income is more than \$50,000 from the 1994 Census database using 14 demographic features such as age, education, marital status, and country of origin. The sensitive attribute is sex. During pre-processing, we remove the sensitive attribute from the set of input features and discard rows with any missing data. The data has 48,842 samples, with 30,527 males and 14,695 females. We use a train-test split ratio of 0.72:0.28.

1138 The 2019) (https://www.kaggle.com/datasets/teertha/ Insurance data (Lantz, 1139 ushealthinsurancedataset) considers predicting the total annual medical expenses of individuals using 5 1140 demographic features from the U.S. Census Bureau, such as BMI, number of children, and age. The sensitive attribute is sex. 1141 During pre-processing, we remove the sensitive attribute from the set of input features and perform normalization. The data 1142 has 1,338 samples with 676 males and 662 females. We use a train-test split ratio 0.8:0.2. 1143

The Crime dataset (Redmond & Baveja, 2002) (https://archive.ics.uci.edu/ml/datasets/ communities+and+crime) considers predicting the number of violent crimes per 100K population using socio-economic information of communities in the U.S. The sensitive attribute is the percentage of people belonging to a particular race in the community. During pre-processing, we drop all the samples with the value of sensitive attribute less than 5% to remove any outliers. We also remove the non-predictive attributes and the sensitive attribute from the set of input features, and normalize all attributes to the standardized range of [0, 1]. The resulting data has 1,112 samples, and we use a train-test split ratio 0.8:0.2.

Implementation details. We use two-layer fully connected networks in all our experiments. For all hidden layers, we use the selu activation function. For the output layer, we use softmax activation for classification and no activation for regression. We use 80 (50) units in the hidden layer and train the network for 30 (200) epochs for classification (regression).

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1175 Figure 11. The performance of Bootstrap-S with $S \in \{3, 9, 27\}$, and Baseline. In the left column, we plot the fraction of 1176 violations of the fairness constraint $\langle \mathbf{a}, \mathbf{b}_{ex} \rangle^2 \leq \varepsilon$ vs. the level of noise added to the sensitive attributes; in the right column, we plot 1177 average MSE vs. the level of noise added to the sensitive attributes.

Empirically, we do not observe any substantial improvement in performance for $S \ge 5$, and thus report for a single S for each dataset. The hyperparameters $\lambda, \lambda_1 \cdots, \lambda_S$ are initialized to 10 (5) for classification (regression). We optimize the model parameters and the hyperparameters $\lambda, \lambda_1 \cdots, \lambda_S$ using different optimizers. For model optimization, we use Adam optimizer with batch size 128 (100), initial learning rate 10^{-3} (10^{-4}), and weight decay 0 (0.01) for classification (regression). For optimizing the hyperparameters, we use SGD optimizer with learning rate 10^{-2} . The batch sizes for Adult, Crime, and Insurance datasets are set to be 128, 100, and 128, respectively. All experiments are implemented in PyTorch using Tesla V100 GPUs with 32 GB memory.

Results of different levels of uncertainty. In Figure 12, we provide results for independence notion of fairness with n = 200 for Adult dataset, $\sigma = 0.25$ for the Crime dataset, and n = 20 for Insurance dataset, respectively, where *n* denotes the number of sensitive attributes kept out of *N* and σ decides the amount of noise added to the sensitive attributes. The observations are consistent to that in Figure 3(a), (b), and (c).



Figure 12. Performance of Bootstrap-S and Baseline for independence for various datasets. Bootstrap-S achieves much better fairness levels compared to Baseline throughout.

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