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SiBBIInGS: Similarity-driven Building-Block Inference using Graphs across States

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Abstract

Time series data across scientific domains are often collected under distinct states (e.g., tasks), wherein latent processes (e.g., biological factors) create complex inter- and intra-state variability. A key approach to capture this complexity is to uncover fundamental interpretable units within the data, Building Blocks (BBs), which modulate their activity and adjust their structure across observations. Existing methods for identifying BBs in multi-way data often overlook inter- vs. intrastate variability, produce uninterpretable components, or do not align with properties of real-world data, such as missing samples and sessions of different durations. Here, we present a framework for Similarity-driven Building Block Inference using Graphs across States (SiBBlInGS). SiB-BlInGS offers a graph-based dictionary learning approach for discovering sparse BBs along with their temporal traces, based on co-activity patterns and inter- vs. intra-state relationships. Moreover, SiBBIInGS captures per-trial temporal variability and controlled cross-state structural BB adaptations, identifies state-specific vs. state-invariant components, and accommodates variability in the number and duration of observed sessions across states. We demonstrate SiBBIINGS's ability to reveal insights into complex biological and medical phenomena through several synthetic and realworld examples. Specifically, we found that SiB-BlInGS recovers meaningful functional neural ensembles underlying Macaque neural recordings and can leverage human EEG data to localize the source of epileptic seizures moments before their onset.

1. Introduction

The analysis of high-dimensional time-series is increasingly important across various scientific disciplines, ranging from neuroscience (Kala et al., 2009; Mudrik et al., 2024) to social sciences (Jerzak et al., 2023) to genetics (Bar-Joseph et al., 2012; Tanvir Ahmed et al., 2023). These data, however, present a daunting challenge in terms of comprehensibility as they are often highly heterogeneous. Specifically, data in many domains are gathered under multiple states (e.g., clinical interventions), while latent factors may introduce variability across trials within states (e.g., internal biological processes that lead to variations in patient responses to treatment). Current analysis methods often struggle to capture the full variability in such multi-state data. Additionally, integrating data from repeated observations (trials) under the same state into a coherent representation is often challenged by missing samples or variable trial duration and sampling rates (Goris et al., 2014; Charles et al., 2018; Duncker & Sahani, 2018). The common practice of within-state trial averaging, for example, obscures important patterns within individual trials.

A promising approach for analyzing multi-state data involves identifying fundamental representational units-Building Blocks (BBs)-whose composition remains similar across states, while their temporal profiles can modulate across trials to capture trial-to-trial variability, both within the same state and across states. These BBs can represent, for instance, neural ensembles in the brain; social groups in diverse contexts; gene clusters under regulatory mechanisms, etc. Identifying these BBs and understanding how they change across states is a key step for recognizing the latent processes underlying the data and providing valuable insights into core commonalities and differences among states. However, uncovering these BBs poses a challenge, as their individual activities or compositions are often unobservable. This challenge is further complicated by potential variations across states, not only in the temporal activity of these BBs, but also in the subtle structural adjustments of the BBs' compositions between states. For example, a neural ensemble may display varying temporal activity between normal (non-seizure) brain activity sessions and during seizures, along with subtle structural adaptations in

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the ensemble composition during seizures (van den Berg &
Friedlander, 2008), e.g., neurons that are not typically part
of the ensemble might become involved during a seizure.

Here, we present SiBBIInGS, a graph-based data driven framework to unravel the complexities of high-dimensional multi-state time-series data, by unveiling its underlying sparse, similarity-driven BBs along with their temporal activity. Our main contributions include:

We develop a novel framework to find interpretable hidden BBs underlying high-dimensional multi-way data while extracting their cross-trial temporal activities and inter-state structural variability, and address real-world challenges unmet by existing methods.

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- We accommodate varying trial conditions, including different time durations, sampling rates, missing samples, and per-state counts, and enable overlapping BB composition.
 - We highlight our method's promise by demonstrating its ability to recover ground-truth components in synthetic data and meaningful latents in several real-world examples.

⁰⁷⁹ 2. Background and Related Work

081 In the case of single-trial analysis, methods for identifying BBs often rely on matrix decomposition including Singular 082 Value Decomposition (SVD) (Kogbetliantz, 1955), Principal 083 Components Analysis (PCA) (Hotelling, 1933), Independent Components Analysis (ICA) (Hyvarinen et al., 2001), 085 or Non-negative Matrix Factorization (NMF) (Lee & Seung, 1999), where sparsity constraints can be added to improve 087 interpretability (e.g., sparse PCA (SPCA)) (Zou et al., 2006). Extending these to the multi-trial setting can be addressed by 089 either concatenating trials end-to-end to create a single wide 090 matrix or by applying these methods individually to each 091 trial. However, this either overlooks the temporal scales 092 of the data (within trial and cross-trial) or ignores shared 093 factors across trials. 094

095 A more suitable extension to multi-trial observations is ten-096 sor methods (TF, e.g., PARAFAC and Tucker decompo-097 sition) (Harshman, 1970; Williams et al., 2018; Mishne 098 et al., 2016; De Lathauwer et al., 2000; Wu et al., 2018), 099 which consider the trials as an additional dimension of the 100 data. However, none of these methods or their combinations with Gaussian processes (GP) (Tillinghast et al., 2020; Xu et al., 2011; Zhe et al., 2016) and dynamic information (e.g., NNDTN and NONFAT (Wang & Zhe, 2022b)) cannot 104 naturally handle variability in trial duration or address state 105 variability as a fourth dimension. Extensions of TF methods 106 seek either identical BBs across states with flexible temporal patterns or fixed temporal traces across states with flexible cross-state BBs. For example, the Shared Response Model 109

(SRM) (Chen et al., 2015) models a multi-subject fMRI model with the same temporal activity across all individuals with different spatial responses between subjects. SRM, however, requires that components be orthogonal, which may not align with biological plausibility. Hyperalignment (HA) (Haxby et al., 2011) addresses a similar setting as SRM by rotating the subjects' time series responses to optimize inter-subject correlation. However, HA does not explicitly reduce the dimension of the feature space.

Other existing methods, such as Dynamic Mode Decomposition (DMD) (Schmid, 2010), model the temporal dynamics explicitly as dynamical systems, however, these methods are tailored for 2D analysis and thus are not designed to simultaneously model multi-state, multi-trial data. State-Space Models (SSMs) (Auger-Méthé et al., 2021) represent another approach to explore time-series data by describing the latent states evolution by a state-transition matrix; however, they do not aim to find sparse interpretable ensembles with cross-trial structural and temporal variability. Other methods include demixed PCA (dPCA) (Brendel et al., 2011), Targeted Dimensionality Reduction (TDR) (Mante et al., 2013) and model-based TDR (mTDR) (Aoi & Pillow, 2018; Aoi et al., 2020). The latter two directly regress rank-1 (TDR) or low-rank (mTDR) components that explicitly target task-relevant variables. However, TDR/mTDR similarly cannot handle trials of varying duration and do not incorporate sparsity in the identified ensembles. dPCA falls short in addressing missing data, different trial durations, and varied sampling rates.

Closer to our approach, dictionary learning (Olshausen & Field, 2004; 1996; Aharon et al., 2006), provides more interpretable representations (Tošić & Frossard, 2011) by learning a feature dictionary where each data point can be linearly reconstructed using only a few of the feature vectors. While traditional dictionary learning treats each data point as independent, recent advances based on re-weighted ℓ_1 (Candes et al., 2008; Garrigues & Olshausen, 2010) can account for spatio-temporal similarities in the sparse feature representations between data points (Garrigues & Olshausen, 2010; Charles & Rozell, 2013; Charles et al., 2016; Zhang & Rao, 2011; Qin et al., 2017; Mishne & Charles, 2019). In particular, re-Weighted ℓ_1 Graph Filtering (RWL1-GF) (Charles et al., 2022) was recently developed for demixing fluorescing components in calcium imaging recordings by correlating the sparse decompositions across a data-driven graph defined by pixel similarity. While RWL1-GF proves the efficacy of graph embeddings in extracting meaningful features, it is constrained to single-trial data and confines its graph construction to a single dimension of the data-the pixel space-overlooking possibly meaningful structures in other dimensions.

Fuzzy clustering (Yang, 1993; Wei et al., 2020; He et al.,

2018) allows data points to exhibit varying degrees of membership in multiple clusters, addressing limitations of methods that restrict data points to a single BB. However, these
approaches focus solely on BB structures rather than their
temporal activities, and do not integrate within-state and
between-state variability information.

116 Nonlinear deep learning models have also been developed 117 to extract latent factors, however, these typically lack inter-118 pretability in mapping back to the feature/sample space, are 119 limited in their ability to produce sparse latent factors, and 120 require large amounts of training data. Variational Autoen-121 coders (VAEs) (e.g., (Xu & Durrett, 2018; Tillinghast & Zhe, 122 2021)) offer recovery of nonlinear latent low-dimensional 123 representations; however, they do not naturally consider the 124 data's temporal structure, and their elements do not directly 125 represent the contribution of individual channels from the 126 input space. Sparse (Ashman et al., 2020; Barello et al., 127 2018) or dynamical (Girin et al., 2020) variants of VAEs do 128 not consider within-vs between state variability of the latent 129 representations. Transformer models, e.g., (Liu et al., 2022) 130 jointly model individual and collective dynamics via an in-131 dividual module for each of several component dynamics 132 and an interaction module that captures pairwise interac-133 tions. However, this necessitates prior knowledge about 134 the system's separation and requires large-scale data due 135 to the encoder-decoder architecture. Recently, for neuronal 136 data analysis, CEBRA (Schneider et al., 2023) incorporates 137 auxiliary labels and temporal information in contrastive op-138 timization, however, it produces a latent state that requires 139 additional interpretation steps to connect to the neuronal 140 space. 141

Notably, all the above approaches are constrained in their
capacity to identify fundamental hidden sparse components
while capturing multi-state, multi-trial variability.

¹⁴⁶¹⁴⁷ 3. Problem Definition and Notations

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148 Consider a system with N channels organized into at most 149 p BBs, with each BB representing a group of channels with 150 shared functionality. These BBs serve as the fundamental 151 constituents of a complex process, however their compo-152 sition is not directly observed nor explicitly known. In 153 particular, let the columns of $A \in \mathbb{R}^{N \times p}$ represent the 154 BBs, such that A_{ij} is the contribution of the *i*-th channel 155 to the *j*-th BB, with $A_{ij} = 0$ indicating that channel *i* 156 does not belong to BB j. We assume that each channel can 157 belong to multiple BBs, and that each BB is sparse (i.e., 158 $\|\mathbf{A}_{j}\|_{0} = K \ll N$ for all $j = 1 \dots p$).

First, we consider a single trial of the system $Y \in \mathbb{R}^{N \times T}$ over T time points. During this trial, each BB exhibits temporal activity denoted by $\Phi \in \mathbb{R}^{T \times p}$, that might reflect current hidden properties of the system, where $\Phi_{t,j}$ is the activity of the *j*-th BB at time *t*. These temporal profiles are assumed to be smooth, bounded (i.e., $||\Phi||_F < \epsilon_1$, for some ϵ_1), and have a low correlation between distinct BBs' activity (i.e., $\rho(\Phi_{:j}, \Phi_{:i}) < \epsilon_2 \quad \forall i \neq j$, for some ϵ_2). In this single trial, our observations, Y, arise from the collective activity of all BBs operating together, $Y = A\Phi^T + \eta$, where η denotes *i.i.d*. Gaussian observation noise. However, the individual composition (A) or activity (Φ) of each BB is unknown.

In the more general setting, we observe a set of M trials, $\{Y_m\}_{m=1}^M$, where the duration of each trial $m = 1 \dots M$ may vary, i.e., $Y_m \in \mathbb{R}^{N \times T_m}$. The BBs (A) remain constant across trials while their corresponding temporal activity ($\Phi_m \in \mathbb{R}^{T_m \times p}$) may vary across trials to capture trial-to-trial variability. The setting we focus on extends beyond a single set of trials; instead, we deal with a collection of D such multi-trial sets, each associated with a known state $d = 1 \dots D$ (Fig. 1A). Across these sets, both the number of trials per set (M_d for each set d), and the durations of the trials, may vary. Thus, the full observation dataset includes the collection of D multi-trial sets, $\{Y_m^1\}_{m=1}^{M_1}, ..., \{Y_m^D\}_{m=1}^{M_D}$, each representing a different state $d = 1 \dots D$, such that $Y_m^d = A^d (\Phi_m^d)^T + \eta_m^d$.

We assume that the BBs' temporal activities ($\{ \Phi_m^d \}$) can vary between trials, both within and between states, and the compositions of the BB ($\{A^d\}$) might present subtle controlled adaptations between, but not within, states. Specifically, we posit that the BBs' dissimilarity between any pair of distinct states d and d' reflects the dissimilarity between those states, such that the distance between A^d and $A^{d'}$ is constrained by $\|A^d - A^{d'}\|_F < \epsilon_3(d, d')$ for some threshold $\epsilon_3(d, d')$ determined by the application. For example, if considering different disease stages as states, we assume that consecutive disease stages are more similar to each other than to a healthy state, such that $\epsilon_3(d_{\text{disease}_1}, d_{\text{disease}_2}) < \epsilon_3(d_{\text{healthy}}, d_{\text{disease}})$. The main challenge that SiBBIInGS addresses is recovering the unknown BBs (A^d) and their temporal activities (Φ_m^d) for all states and trials given only their combined simultaneous activity (Fig. 1A,B).

4. SiBBlInGS

In this section, we present SiBBIInGS—our framework to identify interpretable BBs along with their temporal traces based on shared activation patterns across trials and states. Unlike existing methods, SiBBIInGS identifies BBs in highdimensional data based on temporal similarity without assuming orthogonality, enables BB interdependency or overlap, and can tackle trials of different duration, sampling rates, or count per state (Tab. 2). SiBBIInGS also offers both expert-based supervised and data-driven unsupervised approaches for integrating inter-state similarities, thus pro-



Figure 1. SiBBIInGS Schematic A SiBBIInGS adapts to real-world datasets with varying session durations, sampling rates, and state-specific data by learning interpretable graph-driven hidden patterns and their temporal activity. B SiBBIInGS is based on a perstate-and-trial matrix factorization where the BBs (A^d) are identical across trials and similar across states. C SiBBIInGS controls the BB similarity via data-driven channel graphs ($H^d \in \mathbb{R}^{N \times N}$) and state-similarity graph ($P \in \mathbb{R}^{D \times D}$), which can be either predefined (supervised) or data-driven. D The learning schematic with an exemplary trial for each of the 3 exemplary states. The BBs of each state d (columns of A^d) are constrained with two regularization terms: 1) state-specific λ^d captures similar activity between channels by leveraging the channel-similarity graph H^d , and 2) P, captures BB consistency across states via the state similarity graph. ν controls the relative level of cross-state similarity between BBs, allowing the discovery of both background and state-specific BBs. Higher (lower) ν values promote greater (lesser) consistency of specific BBs across states (e.g. ν_1 v.s ν_5).

viding the flexibility to choose between data-driven or pre-defined approaches based on the specific data structure.

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204 SiBBIInGS is based on an extended dictionary learning-like procedure that alternates between updating the BBs ($\{A^d\}$) 206 and their temporal profiles $(\{ {f \Phi}_m^d \})$ for all states. Critical to our approach is the integration of the non-linear similar-208 ities between both channels and states. We capture these 209 relationships via two graphs (Fig. 1C), one over channels 210 to identify cross-channel regularities, and one over states to 211 promote cross-state similarity in BB structure. Mathemati-212 cally, we formulate the fit $\{\widehat{A}^d\}, \{\widehat{\Phi}^d_m\}$ for all $d = 1 \dots D$ 213

and $m = 1 \dots M_d$ by minimizing the cost function

$$\min_{\{\boldsymbol{A}^d\},\{\boldsymbol{\Phi}_m^d\}} \sum_d^D \left(\sum_m^{M_d} \left[\|\boldsymbol{Y}_m^d - \boldsymbol{A}^d (\boldsymbol{\Phi}_m^d)^T\|_F^2 + \mathcal{R}(\boldsymbol{\Phi}_m^d) \right] + \mathcal{R}(\boldsymbol{A}^d) + \sum_{d' \neq d}^D P_{d,d'} \| (\boldsymbol{A}^d - \boldsymbol{A}^{d'}) \boldsymbol{V} \|_F^2 \right)$$

where the first term is a data fidelity term, and the second term regularizes the BBs' temporal traces. The term $\mathcal{R}(\mathbf{A}^d)$ regularizes each BB to be a sparse group of channels based on shared temporal patterns, and the last term regularizes the BBs' similarity across states (Fig. 1). The use of $\mathbf{V} = \text{diag}(\mathbf{\nu}) \in \mathbb{R}^{p \times p}$, accompanied by the weight vector $\mathbf{\nu} \in \mathbb{R}^p$, allows assigning varying weights to cross-state BBs' similarities to facilitate the discovery of state-invariant vs. state-specific BBs. SiBBIInGS iteratively updates \mathbf{A}^d and



238 Figure 2. Synthetic data results. A Three example time traces identified by SiBBIInGS vs. ground truth traces, projected into the three 239 synthetic states. SiBBInGS recovers both traces that are highly correlated with specific states (e.g., Φ_{10} ; green), as well as traces that 240 exhibit similar activation across states (e.g., Φ_2 ; blue). B Comparison between the identified example BBs and the ground-truth BBs. C 241 Correlation between the example identified time traces and the ground truth (left), and Jaccard index of the identified BBs compared to 242 the ground truth (right). D Comparison between the ground-truth data (top), SiBBIInGS reconstruction (middle), and the residual data 243 (bottom). E Comparison to baseline methods (Sec. 5, App. L). F Performance under noise and random initializations (300 repetitions). 244 Each dot is a model instance. The curve shows the median values, and the shading corresponds to the 25%-75% percentiles. While SiBBIInGS remains robust under varying noise ($\sigma_{\text{signal}}/\sigma_{\text{noise}} > 3$), it experiences a phase transition at a specific noise level, aligning 245 with the dictionary-learning literature (e.g. (Studer & Baraniuk, 2012)). G Performance with increasing levels of missing samples (200 246 repeats). The scattered dots represent model repetitions, the curves depict the median values calculated by rounding to the nearest 5%, and 247 the background shading corresponds to 25%-75% percentiles. 248

 Φ_m^d for each trial and state (Alg. 1 and Fig. 1; computational complexity is described in App. E), as described below:

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Updating A^d : Since we assume that BBs may require 253 subtle state-to-state adaptations but must remain constant 254 within a state, SiBBIInGS demands that the BB matrix (A^d) 255 is shared across same-state trials but undergoes subtle ad-256 justments across states, proportionate to the corresponding 257 states' similarities (P). The update of A^d for each state 258 d, is achieved via an extended re-weighted ℓ_1 graph fil-259 tering with an integration of the channel-similarity graph, 260 $\boldsymbol{H} \in \mathbb{R}^{N \times N}$, in a way that promotes channels with similar 261 temporal activity to be grouped into the same BB. In each updating iteration of A^d , as a pre-calculation step, we first 263 horizontally concatenate the observations from all M_d trials 264 of that d state to receive the matrix $\mathbf{Y}^{d*} \in \mathbb{R}^{N \times (\sum_{m=1}^{M_d} T_m^d)}$, 265 and vertically concatenate the last estimates of the tempo-266 ral traces from all trials of that state to build the matrix 267 $\mathbf{\Phi}^{d*} \in \mathbb{R}^{(\sum_{m=1}^{M_d} T_m^d) \times p}.$ 268 269

We then update each row $n = 1 \dots N$ of \widehat{A}^d $(\widehat{A}_{n:}^d)$ via a re-weighted procedure that alternates between updating

$$\begin{aligned} \boldsymbol{A}_{nj}^{d} \text{ and } \boldsymbol{\lambda}_{n,j}^{d}: \\ \widehat{\boldsymbol{A}}_{n:}^{d} &= \arg\min_{\boldsymbol{A}_{n:}^{d}} \|\boldsymbol{Y}_{n}^{d*} - \boldsymbol{A}_{n:}^{d}(\boldsymbol{\Phi}^{d*})^{T}\|_{2}^{2} + \\ &\sum_{j=1}^{p} \boldsymbol{\lambda}_{n,j}^{d} |\boldsymbol{A}_{n,j}^{d}| + \sum_{d' \neq d} \boldsymbol{P}_{dd'} \| (\boldsymbol{A}_{n:}^{d} - \boldsymbol{A}_{n:}^{d'}) \circ \boldsymbol{\nu}\|_{2}^{2}, \end{aligned}$$
(1)

and

$$\boldsymbol{\lambda}_{n,j}^{d} = \frac{\epsilon}{\beta + |\widehat{\boldsymbol{A}}_{n,j}^{d}| + w_{\text{graph}}|\boldsymbol{H}_{n:}^{d}\widehat{\boldsymbol{A}}_{:j}^{d}|}.$$
 (2)

Above, \circ is element-wise multiplication, and β , ϵ , and w_{graph} are model hyper-parameters. The matrices H^d and P^d are channel and state similarity graphs, described below:

State similarity graph: $P \in \mathbb{R}^{D \times D}$ is a state-similarity graph that determines the effect of the similarity between each pair of states on the regularization of the distance between their BB representations. P can be set manually (supervised P) or in a data-driven way (unsupervised P), thus allowing selection based on specific goals, data type, and knowledge of data labels (Fig. 1C). Each of these two options offers unique benefits: the supervised variant enables explicit regulation of the similarity and the incorporation of human-expert knowledge into the model, whereas the unsupervised variant leverages the data itself to learn similarities
and patterns without preconceived biases.

278 Here, we present the supervised version of P, which is 279 particularly useful when one has prior knowledge or expec-280 tations about quantitative state values that can be leveraged 281 to integrate desired information into the model, while the 282 data-driven approach is presented in Appendix B.2. This 283 supervised version, unlike the data-driven option, assumes 284 that a numerical label L_d , associated with each state d, 285 can provide valuable information for constructing the state-286 similarity graph P (e.g., vector labels that denote x-y co-287 ordinates in a movement task). In this way, the similarity 288 $P_{d,d'}$ between each pair of states (d, d') is calculated based 289 on the distance between the labels $(L_d, L_{d'})$ associated with 290 these states: $P_{d,d'} = \exp\left(-\|L_d - L_{d'}\|_2^2/\sigma_P^2\right)$ where σ_P^2 291 controls how the similarities in labels scale to similarities in 292 BBs. The supervised approach easily extends to both data 293 with identical or different session durations and can also 294 handle categorical states as described in the Appendix B.1.1. 295

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Channel similarity graphs: $H^d \in \mathbb{R}^{N \times N}$ is the 297 channel graph for each state d and is calculated by $H_{i,j}^d = \exp\left(-\|Y_{i:}^{d*} - Y_{j:}^{d*}\|_2^2 / \sigma_H^2\right)$, where σ_H is an hy-298 299 300 perparameter that controls the kernel bandwidth and Y_{i}^{d*} 301 is the horizontally concatenated observations under state d302 described before. To enhance the robustness of $\{H^d\}_{d=1}^D$, we add a post-processing step and utilize the state-graph (P) to re-weigh each H^d along the states dimension: $H^d \leftarrow \sum_{d=1}^{D} P_{d,d'} H^{d'}$. We then retain only the k largest 303 304 305 306 values in each row, setting the rest to zero, symmetrize, and 307 row normalize H^d so that each row sums to one (App. C). 308 This process mitigates the influence of outliers and encour-309 ages the clustering of similarly-behaving channels into the 310 same BB. The advantage of graph-driven re-weighting, com-311 pared to other TF and dictionary learning procedures, is that 312 the updated weighted regularization ($\boldsymbol{\lambda}^d \in \mathbb{R}^{N \times p}$) promotes 313 the grouping (separating) of channels with similar (dissimi-314 lar) activity into the same (different) BBs by integrating the 315 channel similarity graph H^d into the regularization. Specifically, in the last term of the $\lambda_{n,j}^d$'s denominator, for a given state d, a strong (weak) correlation between the temporal neighbors of the n^{th} channel (captured by $H_{n,j}^d$) and the 316 317 318 319 members of the *j*-th BB (\widehat{A}_{ij}^d) results in a decreased (in-320 creased) $\lambda_{n,j}^d$. Consequently, ℓ_1 regularization on $\widehat{A}_{n:}^d$ is 321 reduced (increased)-promoting the inclusion (exclusion) 322 of each channel into BBs that include (exclude) its temporal 323 neighbors. 324

325 After each update of all rows in A^d , each column is normal-326 ized to have a maximum absolute value of 1. In practice, 327 we update A (Eq. (1)) for a random subset of trials in each 328 iteration to improve robustness and computational speed. **Updating** Φ_m^d : The update step over Φ_m^d uses the current estimate of A^d to re-estimate the temporal profile matrix Φ_m^d independently over each state d and trial m. Note that we do not enforce cross-trial similarity in Φ_m^d to allow for flexibility in capturing trial-to-trial variability both within and across states. Thus, for each trial m and state d, $\phi = \Phi_m^d$ is updated by solving:

$$\widehat{\boldsymbol{\phi}} = \arg\min_{\boldsymbol{\phi} \ge 0} \left\| \boldsymbol{Y}_m^d - \boldsymbol{A}^d \boldsymbol{\phi}^T \right\|_F^2 + \gamma_1 \|\boldsymbol{\phi}\|_F^2 +$$
(3)
$$\gamma_2 \|\boldsymbol{\phi} - \widehat{\boldsymbol{\phi}}^{\text{iter}-1}\|_F^2 + \gamma_3 \|\boldsymbol{\phi} - \boldsymbol{\phi}^{t-1}\|_F^2 + \gamma_4 \mathcal{R}_{\text{corr}}(\boldsymbol{\phi})$$

where the first term preserves data fidelity, the second term regularizes excessive activity, the third term encourages continuity across iterations ($\hat{\phi}^{\text{iter}-1}$ is ϕ from the previous iteration), and the fourth term is a diffusion term that promotes temporal consistency of the dictionary (ϕ^{t-1} is ϕ shifted by one time point). $\mathcal{R}_{\text{corr}}(\phi) = \| (\phi^T \phi - \text{diag}(\phi^T \phi)) \circ D \|_{\text{sav}}$ promotes decorrelation of distinct temporal traces, where sav is sum-of-absolute-values and $D \in \mathbb{R}^{p \times p}$ is a normalization matrix with $D_{ij} = \frac{1}{\|\phi_{ij}\|_2 \|\phi_{ij}\|_2}$ (App. D).

5. Experiments

SiBBIInGS recovers ground truth BBs in synthetic data: Synthetic data were generated with D = 3 states, each consisting of a single trial, with p = 10 ground-truth BBs, and N = 100 channels. Each *i*-th BB was generated with a maximum cardinality of $\max_{d,i} ||A_{i,i}^d||_0 = 21$ channels, and on average each channel was associated with 2.1 BBs. While the BBs were designed to be non-orthogonal, we constrained their pairwise correlations to be below a threshold of $\max \rho < 0.6$. The temporal dynamics of the synthetic data were generated by summing 15 trigonometric functions with different frequencies (App. H.2 for details).

SiBBIInGS demonstrated monotonically improving performance during training (Fig. 6A-D), and at convergence was able to successfully recover the underlying BBs in the synthetic data and their temporal traces (Fig 2A-C). In particular, example traces demonstrate a high precision of the recovered temporal traces, with correlation to the ground truth traces being close to one (Fig. 2A, C, 6F). Furthermore, the identified BB components align closely with the ground-truth BBs (Fig. 2B,C), as indicated by high Jaccard index values. We compared SiBBIInGS to existing methods, including Tucker Decomposition, PARAFAC, (S)PCA "global" (a single (S)PCA for all states), (S)PCA "local" ((S)PCA for each state), NONFAT, NNDTN, mTDR, and dPCA, with details in App. L). Notably, SiBBlInGS outperforms existing baselines both in terms of identifying the ground truth BBs and their traces (Fig. 2E, 6F, 8).

SiBBIINGS finds interpretable BBs in Google Trends: We used Google Trends to demonstrate SiBBIINGS' capabil-

SiBBIInGS: Similarity-driven Building-Block Inference using Graphs across States



Figure 3. **Demonstration on Google Trends Data. A** The BBs' temporal traces, as SiBBIInGS found, demonstrate seasonal trends consistent with the terms associated with each BB. **B** Standard deviation of temporal traces over time for the different states align with variability in the states' demographics (Sec. 5). C The BBs SiBBIInGS identified along with their per-state dominancy produce more meaningful clusters than baselines (Fig. 10). States are marked by colors; dot sizes represent the contribution of a term in the BB.



Figure 4. **Identification of Temporal Patterns in Monkey Somatosensory Cortex. A** The reaching out task ((Rodriguez, 2023)). **B** Sparse clusters of neurons representing the identified BBs. **C** Confusion matrix of a multi-class logistic regression model using the inferred temporal traces to predict the state label. **D** The BBs' temporal traces as they vary across states and time. **E** Ratios of within-to-between states temporal correlations for each BB, with $\frac{\rho_{\text{within}}}{\rho_{\text{between}}} > 1$, indicating states distinguishability.

ity in identifying temporal and structural patterns by query-ing search term frequency on Google over time. We used normalized monthly Trends volume of 44 queries (from Jan. 2011 to Oct. 2022) related to various topics, as searched in 8 US states selected for their diverse characteristics (Coulby, 2000) (see pre-processing in App. I.1). The p = 5 BBs iden-tified by SiBBIInGS reveal meaningful clusters of terms, whose time traces convey the temporal evolution of user interests per region (Fig. 3A), while aligning with the sea-sonality of the BBs' components. For instance, the first BB represents college-related terms and shows a gradual annual decrease with periodic activity and a notable deviation dur-ing the COVID pandemic, possibly reflecting factors such as the shift to remote learning (Fig. 3A, 11). The second and

> third BBs, respectively, demonstrate periodic patterns associated with Passover in April (Fig. 12) and winter terms in December. Interestingly, CA, FL, MD, and NY—all states with larger Jewish populations*—show more pronounced peaks of the "Passover" BB activity in April (when Passover is celebrated) compared to the other states (Fig. 3B and 12). The last BB represents COVID-related terms and exhibits temporal patterns with a sharp increase around Jan. 2020, coinciding with the onset of the COVID pandemic in the US. Remarkably, 'Hopkins' exhibits a less pronounced COVIDrelated search peak in MD (blue), where the university and hospital are located, likely attributed to its well-established local presence (Fig. 3C, right). Conversely, other states witnessed a more significant surge in Hopkins-related searches



Figure 5. **Emerging local BBs in Epilepsy**. The recovered BBs under 1) normal activity, 2) activity during the 8 seconds proceedings CPS seizures located around the F8 area, and 3) activity during the seizures. Colors represent different BBs, and the size of the dots corresponds to the contribution of the respective electrode to each BB.

402 at the onset of the COVID outbreak, as Hopkins suddenly 403 garnered increased attention during this period. This empha-404 sizes our model's interpretability and the need to capture 405 similar yet distinct BBs across states. Other methods ap-406 plied with the same number of BBs (p = 5) as used in 407 SiBBIInGS produced less meaningful BBs (Fig. 10).

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408 SiBBIInGS identifies meaningful patterns in brain 409 recordings: We tested SiBBlInGS on neural activity 410 recorded in the somatosensory cortex in a monkey perform-411 ing a reaching task, as described by (Chowdhury & Miller, 412 2022). The data we used include 18 trials under each of the 413 8 hand directions, with each direction corresponding to a 414 unique state. The raw spike data were convolved over time 415 with a Gaussian kernel. When applying SiBBIInGS with a 416 maximum of p = 4 BBs, it identified sparse functional BBs 417 (Fig. 4B) along with meaningful temporal traces (Fig. 4D) 418 that exhibit state-specific patterns. Interestingly, the third 419 BB consistently shows minimal activity across all states, 420 suggesting that it captures background or noise activity. The 421 structure of the identified BBs exhibits subtle yet signifi-422 cant adaptations across states in terms of neuron weights 423 and BB assignments. Furthermore, SiBBIInGS finds neu-424 rons belonging to multiple neural clusters, suggesting their 425 involvement in multiple functions. When examining the 426 temporal correlations of the corresponding BBs within and 427 between states, all BBs exhibited a within/between corre-428 lations ratio > 1 (Fig. 4E, 13C, App. J.4) indicating robust 429 within state trajectories and distinctions between states. Fur-430 thermore, multi-class logistic regression based only on the 431 identified temporal traces accurately predicted the states 432 (Fig. 4C). 433

434 SiBBIInGS discovers emerging BBs preceding seizure:
435 We applied SiBBIInGS to EEG recordings of an epileptic patient from (Handa et al., 2021; Nasreddine, 2021) (App. K).
437 We examined data from an 8-year-old individual who had experienced 5 complex partial seizures (CPS) localized around
439 electrode F8. SiBBIInGS unveiled interpretable and localized EEG activity in the period preceding seizures (Fig. 5), a feat not achieved by other methods. It identified a BB specific to the region around the clinically labeled area (F8) that emerged during the 8 seconds prior to the seizure (Fig. 5, pink circle in SiBBIInGS's middle). Additionally, it found several alterations in the BB composition during the seizure in comparison to the normal activity. E.g., the contribution of T4 to the red BB during normal activity is higher than its contribution during a seizure, while the contribution of T5 to the same BB is larger during a seizure. This underscores the potential of SiBBIInGS in discovering BBs that uniquely emerge under specific states, made possible by the flexibility of ν to support both state-variant and state-invariant BBs.

6. Conclusion

We propose SiBBIInGS for graphs-driven identification of interpretable cross-state BBs with their temporal profiles in multi-way time-series data-providing insights into system structure and variability. Unlike other approaches, SiB-BIInGS supports the discovery of BBs with subtle changes in cross-state structures, allows each channel to belong to a few with varying contributions, and promotes the discovery of both state-invariant and state-specific BBs. We demonstrate SiBBIInGS' capacity to identify functional neural ensembles and discern cross-state variations in web-search data structures, showcasing its promise in additional domains, including, e.g., detecting gene expression clusters in health vs. disease, unveiling local financial patterns based on stock data, and more. Regarding limitations, SiBBIInGS assumes Gaussian statistics, yet Poisson may sometimes be more suitable. Additionally, exploring advanced distance metrics for the graphs construction holds promise for future research. Finally, the identified BBs currently do not consider potential directed connectivity, presenting an exciting future direction.

440 Impact Statement

This paper presents work whose goal is to advance the field of Machine Learning and Time Series Data Analysis. There are several potential societal consequences of our work, none which we feel must be specifically highlighted here.

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Appendix

A. Notations

Throughout the paper, we adopt the following notations: the superscript $()^d$ refers to state d, and the subscript $()_m$ refers to trial m. Specifically, \mathbf{Y}_m^d and Φ_m^d denote the observations and temporal traces of trial m of state d, while \mathbf{A}^d represents the BBs of state d. Additionally, for a general matrix \mathbf{Z} , \mathbf{Z}_i : $(\mathbf{Z}_{:j})$ denotes its *i*-th row (*j*-th column).

Table 1. Notations used in the paper.					
Symbol	Description				
BBs	Building Blocks				
channels	Each feature in the observations, e.g., neurons in recordings				
states	Different "views" of the observations. e.g., different cognitive tasks				
trials/sessions	Repeated observations within state				
p	Number of BBs				
D	Number of states				
M_d	Number of trials for state d				
N	Number of channels				
$Z_{n:}$ (or $Z_{[n,:]}$)	The <i>n</i> -th row of a general matrix \boldsymbol{Z}				
$Z_{:i}$ (or $Z_{[:,i]}$)	The <i>i</i> -th column of a general matrix Z				
L_d	Label of state d (optional, can be a scalar or a vector)				
$oldsymbol{Y}_m^d \in \mathbb{R}^{N imes T_m^d}$	Observation for trial m and state d				
$oldsymbol{A}^{d} \in \mathbb{R}^{N imes p}$	Matrix of BBs for state d.				
$\Phi_m^d \in \mathbb{R}^{T_m^d \times p}$	Matrix of temporal traces for trial m of state d .				
$oldsymbol{P} \in \mathbb{R}^{D imes D}$	States similarity graph				
$\{ oldsymbol{H}^d \}_{d=1}^D, ext{ s.t. } oldsymbol{H}^d \in \mathbb{R}^{N imes N}$	Channel similarity graphs				
$oldsymbol{ u}\in\mathbb{R}^p$	Controls the relative level of cross-state similarity for each BB				
$oldsymbol{V}=diag(oldsymbol{ u})$	A diagonal matrix whose entry in index <i>ii</i> is the <i>i</i> -th entry of ν				
$\epsilon, \beta, w_{\text{graph}}$	Hyperparameters controlling the strength of regularization				
$\gamma_1, \gamma_2, \gamma_3, \gamma_4$	Hyperparameters to regularize $\mathbf{\Phi}_m^d$				
σ_{H}, σ_{P}	Hyperparameters that control the bandwidth of the kernel				
$oldsymbol{\psi}_n^{ij} \in \mathbb{R}^{M_j,M_i}$	Transformation of the data from state i to state j for channel n				

B. Further options for *P* computation

Here, we explore additional approaches for computing the state-similarity graph P. These options take into account factors like data properties, single vs. multi-trial cases, variations in trial duration, and the desired approach (supervised or data-driven).

B.1. Supervised P

In addition to the case presented in the paper, for sequential/ordered states, here we introduce the supervised version designed for categorical or similar-distanced states.

B.1.1. CATEGORICAL OR SIMILAR-DISTANCED STATES

For cases where observation states are represented by categorical labels, and we expect a high degree of similarity between all possible pairs of states (i.e., no pair of labels is closer to each other than to another pair), we can define the state similarity matrix P to ensure uniform values for all entries of distinct states, with larger values assigned to same-state entries located on the matrix diagonal. P is then constructed as

$$\boldsymbol{P} = \boldsymbol{1} \otimes \boldsymbol{1}^T + c\boldsymbol{I},\tag{4}$$

where $\mathbf{1} \otimes \mathbf{1}^T \in \mathbb{R}^{D \times D}$ is a matrix of all ones, $\mathbf{I} \in \mathbb{R}^{D \times D}$ is the identity matrix, and c is a weight that scales the strength of same-state similarity with respect to cross-state similarities.



Figure 6. Synthetic Data Results Robustness - cont. A Model performance under increasing levels of noise, along with random initializations, over the model training iterations. Color: increasing levels of missing samples. Left: Jaccard index between the recovered *A* and the ground true *A*. Middle: Correlation between the recovered *A* and the ground truth *A*. Right: Correlation between the recovered Φ and the ground true Φ . **B** Model performance under random initializations (no noise), over the model training iterations. The blue curve is the median over all repeats. C Model performance under increasing levels of noise only (fixed initializations). **D** Model performance under increasing levels of missing samples, over the model training iterations. E SiBBIInGSPerformance under increasing noise levels. F Comparison to other relevant methods, for each state individually (SiBBIInGSin blue, other methods in pink to red colors).

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Table 2. Assumptions and capabilities comparison between SiBBIInGS and other methods.									
Method	SiBBlInGS	mTDR	PCA	Fast ICA	NMF	GPFA	SRM	HOSVD	PARAFAC
Do not force orthogonality?	V	Х	X	V	V	V	X	Х	V
Sparse?	V	Х	X	X	X	X	X	X	X
Flexible in time across states?	V	Х	Х	X	X	X	X	V	V
Support variations in BB across states?	V	Х	X	X	X	Х	X	Х	X
Used for condition variability?	V	V	Х	X	X	X	X	V	V
Works on tensors?	V	V	X	X	X	Х	V	V	V
Consider both within	V	V	na	na	na	na	X	Х	X
& between states variability?									
Supports state-specific	V	V	na	na	na	na	V	Х	X
emerging components?									
Works on non-consistent	V	Х	na	na	na	na	X	X	X
data duration or									
sampling rates?									
Can prior knowledge	V	V	na	na	na	na	V	X	X
(labels) control state									
similarity?									
Ability to define both	V	V	na	na	na	na	X	X	X
state-specific and									
background components?									
Supports non-negative	V	Х	X	X	V	X	X	X	V
decomposition?									

B.2. Data-Driven P

When prior knowledge about state similarity is uncertain or unavailable, SiBBIInGS also provides an unsupervised, datadriven approach to calculate P based on the distance between data points across states. Here we discuss the four options for constructing the matrix P in a data-driven manner, depending on the structure of the observations.

B.2.1. SINGLE-TRIAL PER-STATE WITH EQUAL-LENGTH ACROSS STATES:

This case refers to the scenario of a single trial for each state $(M_d = 1,$ $\forall d = 1 \dots D$), where all cross-state trials have the same length $(T_1^d = T \quad \forall d = 1...D)$. Here, the similarity graph **P** is constructed as

$$\boldsymbol{P}_{d,d'} = \exp\left(-||\boldsymbol{Y}_1^d - \boldsymbol{Y}_1^{d'}||_F^2 / \sigma_{\boldsymbol{P}}^2\right),\tag{5}$$

where σ_P controls the bandwidth of the kernel.

B.2.2. MULTIPLE TRIALS PER STATE, SAME TRIAL DURATION

In the case where all trials have the same temporal duration, the similarity matrix P is computed by evaluating the distance between the values of each pair of states, considering all trials within each state. For this, we first find the transformation $\psi_n^{ij} \in \mathbb{R}^{M_j \times M_i}$ between the observations of state *i* to the observation of state *j*, by solving the Orthogonal Procrustes problem (Golub & Van Loan, 2013; Gower, 2004). For this, let $Y^{i*} \in \mathbb{R}^{M_i \times (T \times N)}$ be the matrix obtained by vertically concatenating the flattened observations from each trial $(m = 1 \dots M_i)$ of state *i*. Then, the optimal transformation from the observations of state *i* ($\mathbf{Y}^{i*} \in \mathbb{R}^{M_i \times (T \times N)}$) to the observations of state *j* ($\mathbf{Y}^{j*} \in \mathbb{R}^{M_j \times (T \times N)}$) will be

$$\widehat{\boldsymbol{\psi}}^{ij} = \underset{\boldsymbol{\psi}^{ij}}{\operatorname{arg\,min}} \| \boldsymbol{\psi}^{ij} \boldsymbol{Y}^{i*} - \boldsymbol{Y}^{j*} \|_F^2, \tag{6}$$

where this mapping projects the multiple trials of state *i* into the same space as of state *j*, via $\tilde{Y}^{i*} = \hat{\psi}^{ij} Y^{i*}$. The state similarity matrix will thus be

$$\boldsymbol{P}_{ij} = \exp\left(-\|\widetilde{\boldsymbol{Y}}^{i*} - \boldsymbol{Y}^{j*}\|_F^2 / \sigma_p^2\right),\tag{7}$$

for all states $i \neq j = 1 \dots D$, where σ_p controls the kernel bandwidth.

B.2.3. SINGLE-TRIAL PER STATE, SAME DURATION

Further generalization of the state similarity computation requires addressing the case of trials with varying duration. When the observations correspond to the same process and their alignment using dynamic time warping is justifiable, we can replace the Gaussian kernel measure with the Dynamic Time Warping (DTW) distance metric (Berndt & Clifford, 1994). In the case of a single trial for each state, the similarity metric becomes the average DTW distances over all channels,

$$\boldsymbol{P}_{ij} = \exp\left(-\frac{1}{N}\sum_{n=1}^{N}DTW(\boldsymbol{Y}_{n:}^{i}, \boldsymbol{Y}_{n:}^{j})\right).$$
(8)

B.2.4. MULTIPLE TRIALS PER STATE, DIFFERENT DURATION

Similarly, for the multi-trial case we have

$$\boldsymbol{P}_{ij} = \exp\left(-\frac{1}{N}\sum_{n=1}^{N}\left(\frac{1}{M_i}\frac{1}{M_j}\sum_{m_j=1}^{M_j}\sum_{m_i=1}^{M_i}\mathrm{DTW}\left(\left(\boldsymbol{Y}_{m_j}^j\right)_{n:} - \left(\boldsymbol{Y}_{m_i}^i\right)_{n:}\right)\right)\right),\tag{9}$$

where, DTW, is, as before, the Dynamic Time Wrapping (Berndt & Clifford, 1994) operator, applied on the activity of the *n*-th channel in both states. It is crucial to note that this approach operates under the assumption that the trials being compared depict similar processes, and hence aligning them using DTW is a valid assumption.

C. Channel-similarity kernel (H)—generation and processing

The kernel post-processing involves several steps. First, we construct the kernel \widetilde{H}^d for each state $d = 1 \dots D$, as described in the main text. To incorporate similarities between each possible pair of states $d' \neq d$, where $d, d' = 1 \dots D$, we perform a weighted average of each \widetilde{H}^d with the kernels of all other states, using $P_{d:}$ for the weights, as it quantifies the similarity between state d and all other states: $H^d = \sum_{d'=1}^{D} P_{dd'} \widetilde{H}^{d'}$. Then, to promote a more robust algorithm, we only retain the k highest values (i.e., k-Nearest Neighbors; kNN) in each row, while the rest are set to zero. The value of kis a model hyperparameter, and depends on the desired BB size. We then symmetrize each state's kernel by calculating $H^d \leftarrow \frac{1}{2} (H^d + (H^d)^T)$ for all $d = 1 \dots D$. Finally, the kernel is row-normalized so that each row sums to one, as follows: Let Λ^d be a diagonal matrix with elements representing the row sums of H^d , i.e., $\Lambda^d_{ii} = \sum_{n=1}^N H^d_{i,n}$. The final normalized channel similarity kernel is obtained as $H^d_{\text{final}} = (\Lambda^d)^{-1} H^d$.

D. Solving Φ in practice

In Section 4, the model updates the temporal traces dictionary $\phi = \Phi_m^d$ for all $m = 1 \dots M_d$, $d = 1 \dots D$ using an extended least squares for each time point t, i.e.,

$$\widetilde{\phi}_{[t,:]} = \arg\min_{\phi_{[t,:]}} \|\widetilde{Y}^{d}_{m_{[:,t]}} - \widetilde{M}\phi_{[t,:]}\|_{2}^{2},$$
(10)

where $\phi_{[t,:]} \in \mathbb{R}^p$ is the dictionary at time t,

$$\widetilde{\boldsymbol{Y}}_{m_{[:,t]}}^{d} = \begin{bmatrix} \boldsymbol{Y}_{m_{[:,t]}}^{d} \\ [\boldsymbol{0}]_{p \times 1} \\ \gamma_{2} \boldsymbol{\phi}_{[t,:]}^{(iter-1)} + \gamma_{3} \boldsymbol{\phi}_{[(t-1),:]} \end{bmatrix}, \quad \text{and} \quad \widetilde{\boldsymbol{M}} = \begin{bmatrix} \boldsymbol{A}^{d} \\ \gamma_{4} (\frac{[1]_{p \times p}}{p} - \boldsymbol{I}_{p \times p}) \circ \sqrt{D} \\ (\gamma_{1} + \gamma_{2} + \gamma_{3}) \boldsymbol{I}_{p \times p} \end{bmatrix},$$

with all parameters being the same as those defined in Section 4 of the main text and \circ denotes element-wise multiplication. Here, $[0]_{p \times 1} \in \mathbb{R}^{p \times 1}$ represents a column vector of zeros, $[1]_{p \times p}$ represents a square matrix of ones with dimensions $p \times p$, and $\mathbf{Y}_{m_{[:,t]}}^d \in \mathbb{R}^N$ denotes the measurement in the *m*-th trial of state *d* at time *t*.

E. Model Complexity

SiBBIInGS relies on 4 main computational steps:

Channel Graph Construction: This operation, performed once for all N channels of every state $d = 1 \dots D$, generates a channel graph $H^d \in \mathbb{R}^{N \times N}$ for each state $d \in [1, D]$ by concatenating within-state trials $1 \dots M_d$ horizontally, resulting in a $N \times \sum_{m=1}^{M_d} T_m^d$ matrix. For simplicity, let $\tilde{T} = \sum_{m=1}^{M_d} T_m^d$. The computational complexity of calculating the pairwise similarities of this concatenated matrix for all D states is thus $\mathcal{O}\left(D\tilde{T}^2N(N-1)\right)$.

For the k-threshold step (B.2.1), that involves keeping only the *k* largest values in each row while setting the other values to zero—the complexity will be $\mathcal{O}\left(\tilde{T}\log k\right)$ per row for a total computational complexity of $\mathcal{O}\left(DN\tilde{T}\log k\right)$ for *N* rows and *D* states.

889 State Graph Construction: This is a one-time operation that involves calculating the pairwise similarities between each 890 pair of states. For simplicity, if we assume the case of user-defined scalar labels, and as in this case there are D states (and 891 accordingly D labels), the computation includes D(D-1)/2 pairwise distances for $O(D^2)$.

892 **BB Inference (Eq.** (1)): This iterative step involves per-channel re-weighted ℓ_1 optimization. If the computational 893 complexity of a weighted ℓ_1 is denoted as C, then the computational complexity of the re-Weighted ℓ_1 Graph Filtering is 894 NLC + LNk, where N is the number of channels, L is the number of iterations for the RWLF procedure, and k is the 895 number of nearest neighbors in the graph. For the last term in Eq. (1), there are p^2 multiplicative operations involving 896 the vector ν and the difference in BBs, arising from the ℓ_2^2 norm. Additionally, there is an additional multiplication step 897 involving $P_{dd'}$. For each state d, this calculation repeats itself D-1 times (for all $d' \neq d$). This process is carried out for 898 every $d = 1 \dots D$. In total, these multiplicative operations sum up to $(p^2 + 1) D(D - 1)$, resulting in a computational 899 complexity of $\mathcal{O}(D^2p^2)$. 900

901 **Optimization for** ϕ : This step refers to the least-squares problem presented in Eq. (10) in Appendix D. If a non-902 negative constraint is applied, SiBBIInGS uses scipy's "nnls" for solving $\tilde{\phi}_{[t,:]} = \arg \min_{\phi_{[t,:]}} \|\tilde{Y}_{m_{[:,t]}}^d - \tilde{M}\phi_{[t,:]}\|_2^2$, 903 where $\tilde{Y}_m^d \in R^{(N+2p) \times T_m^d}$ and $M \in R^{(N+2p) \times p}$. This results in complexity of $\mathcal{O}\left(p(N+2p)^2F\max(T_m^d)\right)$, where F is 904 the number of nnls iterations. Without non-negativity constraint, this problem is a least squares problem with a complexity 905 of $\mathcal{O}\left(\max(T_m^d)p^2(2p+N)\right)$. Potential complexity reduction options include parallelizing RWL1 optimizations per 907 channel, using efficient kNN or approximate kNN search for constructing kNN graphs instead of full graphs, and employing 908 dimensionality reduction techniques to expedite nearest neighbor searches.

F. Data and Code Availability

The code employed in this study will be made publicly available on GitHub upon publication and is included in the supplementary material of the submission. The data used in this study are publicly available and cited within the paper.

916 Inputs 917 $\{Y_m^1\}_{m=1}^{M_1}, ... \{Y_m^D\}_{m=1}^{M_D} \{\text{Observations}\}$ 918 $\beta, \epsilon, \gamma_1, \gamma_2, \gamma_3, \gamma_4, w_{graph}, \sigma_p, \sigma_H \{\text{Hyperparameters}\}$ 919 **Initialization and pre-Calculations** 920 $\{ \boldsymbol{A}^{d} \}, \{ \boldsymbol{\Phi}_{m}^{d} \}_{m=1}^{M_{d}}, \forall d = 1 \dots D \{ \text{Initialize BBs\&traces} \}$ $\boldsymbol{P} \in \mathbb{R}^{D \times D} \{ \text{Calculate state-similarity graph} \}$ $\boldsymbol{H}^{d} \in \mathbb{R}^{N \times N}, \forall d = 1 \dots D \{ \text{Calculate channels graphs} \}$ 921 922 923 repeat 924 for $d = 1 \dots D$ do 925 Select a random batch of trials from state d926 Update A^d and λ^d {via Eq. (1) and Eq. (2)} for $m = 1 \dots M_d$ do Update Φ_m^d {via Eq. (3)} 928 929 end for 930 end for 931 until BBs and traces of all states converged 932 933

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Figure 7. **BBs identified by different methods**. BBs identified by SiBBIInGS are compared with those from other methods, including PARAFAC, Tucker, PCA (global and local), Sparse PCA (global and local), demixed PCA (Kobak et al., 2016), mTDR (Aoi & Pillow, 2018), and Gaussian-process-based methods (Wang & Zhe, 2022b). The identified BBs were reordered to best match the ground truth BBs' temporal traces through maximum correlations. A subsequent hard-thresholding step was applied to achieve sparsity, aligning with the sparsity level with of the ground truth components. The BBs were normalized to have an absolute sum of 1 each for visualization purposes.

G. General Experimental Details

All experiments and code were developed and executed using Python version 3.10.4 and are compatible with standard desktop machines.

H. Synthetic Data—Additional Information

979 980 H.1. Synthetic Generation Details

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We initiated the synthetic data generation process by setting the number of channels to N = 100 and the maximum number of BBs to p = 10. We further defined the number of states as D = 3 and determined the number of time points in each observation to be $T^d = 300$, where d represents the state index (here $d \in \{1, 2, 3\}$). We defined the number of trials for each state as one, i.e., $M_d = 1$ for d = 1, 2, 3.

986 We first initialized a "general" BB matrix (A) as the initial structure, which will later undergo minor modifications for each 987 state.

For each state d, we generated the time-traces Φ^d via a linear combination of 15 trigonometric signals, such that the



Figure 8. Correlations between BBs identified by different methods and ground truth BBs for each state and BB number.

temporal trace of the *j*-th BB was defined as $\Phi_{:j}^d = \sum_{i=1}^{15} c_i f_i (\text{freq}_i * x)$ where *x* is an array of T = 300 time points (x = [1, ..., 300]), freq_i is a random scaling factor sampled uniformly on [0, 5], *f* refers to a random choice between the sine and cosine functions (with probability 1/2 for each), and the sign (c_i) was flipped (+1 or -1) with a probability of 1/2.

1031 During the data generation process, we incorporated checks and updates to A and Φ to ensure that the BBs and their 1032 corresponding time traces are neither overly correlated nor orthogonal, are not a simple function of the states labels, and 1033 that different BBs exhibit comparable levels of contributions. This iterative process involving the checks persisted until no 1034 further modifications were required.

The first check aimed to ensure that the temporal traces of at least two BBs across all states were not strongly correlated with the state label vector ([1, 2, 3]) at each time point. Specifically, we examined whether the temporal traces of a *j*-th BB across all states ($\Phi_{tj}^1, \Phi_{tj}^2, \Phi_{tj}^3$) exhibited high correlation with the state label vector at each time point. This check was important to avoid an oversimplification of the problem by preventing the temporal traces from being solely influenced by the state labels. To perform this check, we calculated the average correlation between the temporal traces and the state labels ([1, 2, 3]) at each time point. If the average correlation over time exceeded a predetermined threshold of 0.6, we introduced additional variability in the time traces of the BB that exhibited a high correlation with the labels. This was achieved by adding five randomly generated trigonometric functions to the corresponding BB. These additional functions were generated in the same manner as the original data (with $\Phi_{:j}^d = \sum_{i=1}^5 c_i f_i(\text{freq}_i \cdot x)$). 1045 The second check ensured that the time traces were not highly correlated with each other and effectively represented separate 1046 functions. If the correlation coefficient between any pair of temporal traces of different BBs exceeded a threshold of $\rho = 0.6$, 1047 the correlated traces were perturbed by adding zero-mean Gaussian noise with a standard deviation of $\sigma = 0.02$.

Next, we ensured that the BBs represented distinct components by verifying that they were not highly correlated with each other. Specifically, if the correlation coefficient between a pair of BBs (A_{ij} , A_{ii} for j, i = 1...10) within a state exceeded the threshold $\rho = 0.6$, each BB in the highly-correlated pair was randomly permuted to ensure their distinctiveness.

1051 the inreshold $\rho = 0.6$, each BB in the highly-correlated pair was randomly permuted to ensure their distinctiveness.

1052 To prevent any hierarchical distinction or disparity in BB contributions and distinguish our approach from order-based 1053 methods like PCA or SVD, we evaluated each BB's contribution by measuring the increase in error when exclusively using 1054 that BB for reconstruction. Specifically, for the *j*-th BB, we calculated its contribution as contribution_{*j*} = $-||Y - A_{:j} \otimes$ 1055 $\Phi_{:j}||_F$, where \otimes denotes the outer product. Then, we compared the contributions between every pair of BBs within the 1056 same state. If the contribution difference between any pair of BBs exceeded a predetermined threshold of 10, both BBs in 1057 the pair were perturbed with random normal noise. Subsequently, a hard-thresholding operation was applied to ensure that 1058 the desired cardinality was maintained.

To introduce slight variability also in the BBs' structures across states (in addition to the temporal variability), the general basis matrix A underwent modifications for each of the states. In each state and for each BB, a random selection of to to 2 non-zero elements from the corresponding BB in the original A matrix were set to zero, effectively introducing missing channels as differences between states, such that A^d is the updated A modified for state d. Finally, the data was reconstructed using $Y^d = A^d (\Phi^d)^T$ for each state d = 1, 2, 3.

1065 1066 H.2. Experimental details to the Synthetic data

We applied SiBBLInGS to the synthetic data with p = 10 components and a maximum number of 10^3 iterations, while in practice about 50 iterations were enough to converge (see Fig. 6). The parameters for the λ update in Eq. (2) were $\epsilon = 0.01$, $\beta = 0.09$, and $w_{\text{graph}} = 1$. For the regularization of Φ (Eq. (3)), the parameters used were $\gamma_1 = 0.1$, $\gamma_2 = 0.1$, $\gamma_3 = 0$, and $\gamma_4 = 0.0001$. ν was set to be a vector of ones with length p = 10. The number of repeats for the update of A within an iteration, for each state, was set to 2. The number of neighbors used in the channel graph reconstruction (H^d) was k = 25. The python scikit-learn's (Pedregosa et al., 2011) LASSO solver was used for updating A in each iteration. This synthetic demonstration used the supervised case for building P, where P was defined assuming similar similarity levels between each pair of states, by defining $P = \mathbf{1} \otimes \mathbf{1}^T \in \mathbb{R}^{3\times 3}$ (the case described in App. B.1.1, with c = 1).

1076 H.3. Jaccard index calculation

1078 In Figure 1C, we computed the Jaccard similarity index between the identified BBs by SiBBIInGS and the ground truth BBs. 1079 To obtain this measure, we first rearranged the BBs based on the correlation of their temporal traces with the ground truth 1080 traces (since the method is invariant to the order of the BBs). Then, we nullified the 15 lower percentiles of the \hat{A} matrix, 1081 which correspond to values very close to zero. Finally, we compared the modified identified BBs to the ground truth BBs 1082 using the "jaccard_score" function from the sklearn library (Pedregosa et al., 2011).

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1084 **I. Google Trends—Additional Information**

1086 I.1. Trends data acquisition and pre-processing

The acquisition and pre-processing of Google Trends data involved manually downloading the data from April 1, 2010, to November 27, 2022, for each of the selected states: California (CA), Maryland (MD), Michigan (MI), New York (NY), Illinois (IL), Louisiana (LA), Florida (FL), and Washington (WA), directly from the Google Trends platform (Google Trends, Accessed 11 November 2022). The full list of terms (including their clusters as SiBBIInGS recovered) is presented in Figure 9. To ensure coverage of search patterns, the data was downloaded by examining each query in all capitalization formats, including uppercase, lowercase, and mixed case.

1094The data (in CSV format) was processed using the "pandas" library in Python (pandas development team, 2020; Wes1095McKinney, 2010) and keeping only the relevant information from January 2011 to October 2022, inclusively. We conducted1096a verification to ensure the absence of NaN (null) values for each term in every selected state. To pre-process each term, we

1097 implemented a two-step normalization procedure. First, the values within the chosen date range were scaled to a maximum 1098 value of 100. This step ensured that the magnitude of each term's fluctuations remained within a consistent range. Next,

Sidding Similarity-united bunding-block interence using Graphs across State	SiBBIInGS: Similarity-driv	en Building-Block Infere	nce using Graphs across States
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	СА	FL	IL	LA	MD	MI	NY	WA
BB 1	Berkeley, Campus, College, Harvard, Phd, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton	Admissions, Berkeley, Campus, College, Harvard, Princeton	Admissions, Campus, College, Harvard, Phd, Princeton
ВВ 2	Afikomen, Chametz, Charoset, Haggadah, Pesach, Seder	Afikomen, Chametz, Charoset, Haggadah, Passover, Pesach	Chametz, Charoset, Haggadah, Passover, Pesach, Seder	Chametz, Charoset, Haggadah, Passover, Pesach, Seder	Afikomen, Chametz, Charoset, Haggadah, Passover, Pesach	Chametz, Charoset, Haggadah, Passover, Pesach, Seder	Afikomen, Chametz, Charoset, Haggadah, Pesach, Seder	Chametz, Charoset, Haggadah, Passover, Pesach, Seder
вв 3	Auld lang syne, Chicken soup, Decorations, Depression, Gpa, Sweets	Auld lang syne, Champagne, Chicken soup, Depression, Gpa, Sweets	Auld lang syne, Chicken soup, Decorations, Depression, Gpa, Sweets	Auld lang syne, Champagne, Chicken soup, Decorations, Gpa, Sweets	Auld lang syne, Champagne, Chicken soup, Decorations, Gpa, Sweets	Auld lang syne, Champagne, Chicken soup, Decorations, Gpa, Sweets	Auld lang syne, Champagne, Chicken soup, Countdown, Decorations, Sweets	Auld lang syne, Champagne, Chicken soup, Decorations, Depression, Sweets
вв 4	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater	Elf, Gift, New years eve, Poinsettia, Ugly sweater
ВВ 5	Cdc, Hopkins, Kippur, N95, Quarantine, Zoom	Cdc, Hopkins, Mit, N95, Quarantine, Zoom	Cdc, Hopkins, N95, Quarantine, Zoom	Cdc, Hopkins, Mit, N95, Quarantine, Zoom	Cdc, Hopkins, N95, Quarantine, Zoom	Cdc, Hopkins, N95, Quarantine, Zoom	Cdc, Hopkins, N95, Quarantine, Zoom	Cdc, Hopkins, Mit, N95, Quarantine, Zoom

Figure 9. Table of clustered words for the Google Trends experiment

	SiBBIInGS	PCA Local	PCA Global	PARAFAC	Tucker
	Cdc, Hopkins, N95,		Hopkins, N95, New years eve, Quarantine, Ugly		
BB 1	Quarantine, Zoom	Hopkins, N95, Quarantine, Zoom	sweater, Zoom		
	Afikomen, Chametz,	Afikomen, Ball drop, Charoset, Elf,	Afikomen, Ball drop, Berkeley, Chametz, Depression,	Admissions, Afikomen, Berkeley, Cdc, Chametz, Charoset,	Elf, Hopkins, N95, New years eve, Poinsettia,
	Charoset, Haggadah,	Gift, Haggadah, Memorial	Gpa, Haggadah, Harvard, Memorial, N95, New years	Decorations, Depression, Gpa, Haggadah, Harvard, Instacart,	Quarantine, Santa
BB 2	Pesach, Seder	N95, Pesach, Ugly sweater, Zoom	eve, Passover, Pesach, Seder, Spirit	Labor, Matzo ball, Passover, Pesach, Princeton, Seder, Spirit	Ugly sweater, Zoom
	Auld lang syne, Chicken				
	soup, Decorations,	Charoset, Elf, Memorial, New	Cdc, Chametz, Charoset, Haggadah, N95, Passover,	Admissions, Ball drop, Cdc, Countdown, Hopkins, Instacart,	Cdc, Chametz, Charoset, Haggadah, N95, Passover,
BB 3	Depression, Gpa, Sweets	years eve, Ugly sweater	Pesach, Quarantine, Seder, Zoom	N95, New years eve, Quarantine, Zoom	Pesach, Quarantine, Seder, Zoom
				Auld lang syne, Champagne, Decorations, Elf, Gift, Hopkins,	
	Elf, Gift, New years eve,		Afikomen, Auld lang syne, Ball drop, N95, New years	Labor, Memorial, N95, New years eve, Poinsettia, Santa, Ugly	Campus, Charoset, Elf, Labor, Memorial, New years
BB 4	Poinsettia, Ugly sweater	Cdc, N95, Quarantine, Zoom	eve, Ugly sweater	sweater	eve, Spirit, Ugly sweater
	Berkeley, Campus, College,	Afikomen, Charoset, Memorial,			Auld lang syne, Ball drop, Labor, N95, Memorial, New
BB 5	Harvard, Phd, Princeton	Zoom	Charoset, Elf, Labor, New years eve, Ugly sweater	Auld lang syne, Ball drop, N95, New years eve	years eve, Ugly sweater

Figure 10. **Comparison of The Google Trends Results to Other Methods with 5 BBs for CA:** Comparison to other methods, each applied with p = 5 BBs, yielded less interpretable BBs. For example, SiBBlInGS discerns theme-specific BBs (e.g., 'Covid' and 'University'), while other methods produce more blended compositions. Empty cells for PARAFAC and Tucker indicate that those BBs remained empty.

the values for each term were divided by the sum of values across all dates and then multiplied by 100, resulting in an adjusted scale where the area under the curve for each term equaled 100. This normalization procedure accounted for potential variations in the frequency and magnitude of term occurrences, enabling fair comparisons across different terms. By applying these pre-processing steps, we aimed to mitigate the influence of isolated spikes or localized peaks that could distort the overall patterns and trends observed in the data. Since the focus of this processing was on assessing the relative contribution of a term within a BB rather than comparing the overall amplitude and mean of the term across states, factors such as population size and other characteristics of each state were not taken into consideration.

I.2. Experimental details for Google trends

We ran the Trends experiment with p = 5 BBs, and applied non-negativity constraints to both the BB components and their temporal traces. The λ 's parameters in Equation (1) included $\epsilon = 9.2$, $\beta = 0.01$, and $w_{graph} = 35$. For the regularization of Φ in Equation (3), we used the parameters $\gamma_1 = 0$, $\gamma_2 = 0$, $\gamma_3 = 0.05$, $\gamma_4 = 0.55$. The trends example used the data-driven version for studying P, and we set ν to be a vector of ones with length p = 5.

During each iteration, A underwent two updates within each state d. The number of neighbors we used in the channel graph reconstruction was k = 4. We used the PyLops package in Python, along with the SPGL1 solver (Ravasi & Vasconcelos, 2020) to update A in each iteration. With respect to SPGL1 parameters (as described in (Ravasi & Vasconcelos, 2020)), we set the initial value of the parameter τ to 0.12, and a multiplicative decay factor of 0.999 was applied to it at each iteration. We note here that SPGL1 solves a Lagrangian variation of the original Lasso problem, where, i.e., it bounds the ℓ_1 norm of the selected BB to be smaller than τ , rather than adding the ℓ_1 regularization to the cost (van den Berg & Friedlander, 2008; Ravasi & Vasconcelos, 2020).



1171 Figure 11. Temporal traces of one of the BBs SiBBIInGS identified, which includes college admission terms, show bi-yearly peaks around March and October, aligning with key milestones in the US college admissions process. Additionally, a decrease in online interest in the college BB is observed during the COVID-19 pandemic. 1174

1176 I.3. Temporal traces of college BB

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1177 The temporal traces of the BBs that relate to college admission, as identified by SiBBIInGS, exhibit distinct bi-yearly peaks, 1178 with notable increases in activity around March and October, along with a clear decrease between March to next October 1179 (Fig. 11). These peaks align with key periods in the US college admissions cycle, including application submission and 1180 admission decision releases. Particularly, around the end of March, many colleges and universities in the US release their 1181 regular admission decisions, prompting increased population interest. Similarly, October marks the time when prospective 1182 students typically start showing increased interest in applying to colleges, as many colleges have early application deadlines 1183 that fall in late October or early November. The bi-yearly peaks pattern in March and October thus reflects the concentrated 1184 periods of activity and anticipation within the college admissions process. External factors such as the COVID-19 pandemic 1185 can also influence the timing and dynamics of the college admissions process, as we observe by the decrease in the college 1186 BB activity during the pandemic period (Fig. 11). 1187

1188 I.4. Temporal traces of "Passover" BB 1189

1190 SiBBINGS identified a "Passover" BB, characterized by temporal traces that show a clear alignment with the timing of 1191 Passover, which usually occurs around April. The time traces demonstrate a prominent peak in states with higher Jewish 1192 population percentages, like CA, FL, and NY, as presented by the average peak value in Figure 12B plotted for the different 1193 states. The peaks detection (in Fig. 12) was done using scipy's (Virtanen et al., 2020) "find_peaks" function with a threshold 1194 of 4. 1195

1196 J. Neural Data—Additional Information

1198 J.1. Neural Data Pre-Processing 1199

In this experiment we used the neural data collected from Brodmann's area 2 of the somatosensory cortex in a monkey 1200 performing a reaching-out movement experiment from Chowdhury et al. (Chowdhury & Miller, 2022; Chowdhury et al., 2020). While the original dataset includes data both under perturbed and unperturbed conditions, here, for simplicity, we used only unperturbed trials. We followed the processing instructions provided by Neural Latents Benchmark (Pei et al., 2021) to extract the neural information and align the trials. The original neural data consisted of spike indicators per neuron, 1204 which we further processed to approximate spike rates by convolving them with a Gaussian kernel. 1205

1206 For each of the 8 angles, we randomly selected 18 trials, resulting in a total of 144 data matrices. The states were defined as the angles, and for learning the supervised P, we used as labels the x-y coordinates of each angle in a circle with a radius of 1 (i.e., sine and cosine projections). 1209



1226 Figure 12. Temporal trace of Passover BB. The Passover's BB patterns, as SiBBIInGS found, show an alignment with the percentage of Jewish population in different states. A Temporal traces of the Passover BB for each state. Vertical black lines indicate the month of April, 1228 when Passover is usually celebrated. **B** The mean and standard error of peak values for each state.

J.2. Experimental details for the neural data experiment 1231

1232 We ran SiBBLINGS on the reaching-out dataset with p = 4 BBs. The λ 's parameters used were $\epsilon = 2.1, \beta = 0.03$, and $w_{\text{graph}} = 10.1$. For the regularization of Φ we used: $\gamma_1 = 0.001$, $\gamma_2 = 0.001$, $\gamma_3 = 0.1$, and $\gamma_4 = 0.3$ and we set ν to be a 1234 vector of length p = 4 with $\nu_1 = 0.8$ (to allow more flexibility in the first BB), and $\nu_k = 1$ for k = 2, 3, 4. For the neural 1235 data, we used the supervised version of P, where the x - y coordinates are used as the labels for calculating P. During 1236 each iteration, A underwent two updates within each state. We chose k = 7 neighbors for the channel graph reconstruction, and used Python scikit-learn's (Pedregosa et al., 2011) LASSO solver for the update of A.

1239 J.3. State prediction using temporal traces 1240

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1241 We used the identified temporal traces Φ to predict the state (hand direction). The dimensionality of each state's temporal 1242 activity $\mathbf{\Phi}^d$ was reduced to a vector of length $p \times 4 = 16$ using PCA with 4 components. A k-fold cross-validation 1243 classification approach with k = 4 folds was used in a multi-class logistic regression model with multinomial loss (trained 1244 on 3 folds and used to predict the labels of the remaining fold). This process was repeated for each fold, and the results were 1245 averaged. The confusion matrix and accuracy scores for each state (angle), are shown in Figure 4C and in Figure 13F. 1246

1247 J.4. Computation of $\rho_{\text{within/between}}$ 1248

To compute the correlation for the "within" state case, a random bootstrap approach was used, such that, for each state, we 1249 randomly selected 100 combinations of temporal trace pairs corresponding to the same BB from different random trials 1250 within the same state. We computed the correlations between these temporal trace pairs, and averaged the result over all 100 1251 bootstrapped pairs to obtain the average correlation. 1252

1253 Similarly, for the "between" states case, we repeated this procedure for pairs of trials from distinct states. Particularly, we 1254 selected 100 random bootstrapped combinations of pairs of the same BB from trials of different states. In Figure 13C, the 1255 average correlations are shown for each BB. The ratio depicted in Figure 4E represents the ratio between the averages of the 1256 "within" and "between" state correlations.

K. Epilepsy—Additional Information 1259

1260 K.1. Data Characteristics and Pre-processing for SiBBIInGS Analysis 1261

The Epilepsy EEG experiment in this paper is based on data kindly shared publicly in (Handa et al., 2021). 1262

The data consist of EEG recordings obtained from six patients diagnosed with focal epilepsy, who were undergoing 1264

1265 presurgical evaluation. As part of this evaluation, patients temporarily discontinued their anti-seizure medications to 1266 facilitate the recording of habitual seizures. The data collection period spanned from January 2014 to July 2015. These 1267 seizures manifest different patients, seizure types, ictal onset zones, and durations.

The EEG data, as described by (Handa et al., 2021), were recorded using a standard 21 scalp electrodes setup, following the 10-20 electrode system, with signals sampled at a rate of 500 Hz. To enhance data quality, all channels underwent bandpass filtering, with a frequency ranging from 1/1.6 Hz to 70 Hz. Certain channels, including Cz and Pz, were excluded from some recordings due to artifact constraints.

1273 Here, we focused on the EEG data from an 8-year-old male patient. This patient experienced five recorded complex partial 1274 seizures (CPS) in the vicinity of electrode F8. The EEG data for this patient includes both an interictal segment during 1275 which no seizures are recorded and 5 ictal segments representing seizures.

To prepare the data for compatibility with the input structure of SiBBIInGS, we divided the epileptic seizure data into non-overlapping batches, with a maximum of 8 batches extracted from each seizure. Each batch had a duration of 2000 time points, equivalent to 4 seconds. This process resulted in 4 seizures with 8 batches each and one seizure with 7 batches due to its shorter duration.

For each seizure, we also included data from the 8 seconds preceding the marked clinical identification of the seizure. This amounted to 2 additional 2000-long batches (each corresponding to 4 seconds) before each seizure event.

Regarding normal activity data, we randomly selected 40 batches, each spanning 4 seconds (2000 time points), from various time intervals that did not overlap with any seizure activity or the 8-second pre-seizure period.

1286 In total, we had 40 batches of normal activity, 39 batches of seizure activity, and 10 batches of pre-seizure data.

We ran SiBBIInGS on this data with p = 7 BBs. For the state-similarity graph (P), we adopted a supervised approach to distinguish between seizure and non-seizure states, as detailed in the categorical case in Appendix B.1.1, where we assigned a strong similarity value constraint to same-state trials and lower similarity values to different-state trials.

We also leverage this example to underscore the significance of the parameter ν in the model's ability to discover networks that emerge specifically under certain states as opposed to background networks. In this context, we defined here $\nu = [1, 1, 1, 1, 1, 1, 0]$ such that the similarity levels of the 1st to 6th BBs are determined by the relevant values in P, while the last BB's similarity is allowed to vary between states.

During the training of SiBBIInGS on this data, we adopted a training strategy where 8 random batches were selected in each iteration to ensure that the model was exposed to an equal number of trials from each state during each iteration and enhancing its robustness.

12991300 L. Additional Details about the Baselines

Initial Extraction of BBs from each method: To compare SiBBIInGS with other methods (as presented in Fig. 10, 2, 5), we took the following approach. For PCA global, we applied PCA on the entire dataset after horizontally concatenating the time axis using the sklearn (Pedregosa et al., 2011) implementation, specifying the number of Principal Components (PCs) to match the number of BBs in the ground truth data (p = 10). These PCs were then treated as the BBs. In the case of PCA local, we followed a similar procedure. However, we ran PCA individually for each state.

For Sparse PCA global (SPCA global) and Sparse PCA local (SPCA local) we used its sklearn implementation, while tuning the sparsity level on the BBs (the α parameter) to match the sparsity level of the ground truth data. Similar to the regular PCA, SPCA global refers to applying a single SPCA on the entire dataset with p = 10 components, while and Sparse PCA local (SPCA local) refers to applying SPCA to the observations of each state.

For dPCA, we used the Python implementation offered here (Kobak et al., 2016) with k = 3 states and p = 10 components. We chose to protect the time axes within each trial against shuffling ("dpca.protect = ['t']") and extracted the temporal traces using the stimulus component ("s" key) of the trained model.

1315 For Tucker and PARAFAC, we utilized the Tensorly library (Kossaifi et al., 2021) with a rank set to p = 10 (the number of

1316 BBs allowed by SiBBIInGS). We interpreted the BBs as the first factor (factors[0] in Tensorly output), and we considered

1317 the temporal traces as the second factor (factors[1] in Tensorly output) while multiplying them by the corresponding weights

1318 from the state factor (third factor, factors[2]) to enable cross-state flexibility to these temporal traces.

SiBBIInGS: Similarity-driven Building-Block Inference using Graphs across States

- For mTDR (Aoi & Pillow, 2018; Aoi et al., 2020), we first note that this method focuses on a slightly different problem 1321 than SiBBIInGS, specifically tailored for cases where multiple conditions influence each trial simultaneously. Hence, in 1322 our comparative analysis, we first adapted mTDR to be comparable with SiBBIInGS by applying the following processing 1323 steps. Given that the synthetic example (Fig.2) involves categorical rather than ordinary sequential states, we changed 1324 the categorical states to dummy variables using one-hot encoding before running mTDR. We then ran mTDR on the 1325 concatenation of all trials to obtain the temporal basis matrices (S as denoted in (Aoi & Pillow, 2018)) and their neuron-1326 specific weights W. We recalculated the optimal coefficients based on the identified S and W to minimize the Mean 1327 Squared Error (MSE) in reconstructing each trial, and then obtained the state-specific temporal activity through optimal 1328 re-weighting of S. We used the reweighed optimal S to compare it with Φ from our paper's notation, while mTDR's W 1329 served as the structural matrices for comparison with A in our notation. 1330 For NONFAT (Wang & Zhe, 2022b), we used the code shared by the authors at (Wang & Zhe, 2022a). The model was 1331 executed with the same parameters as specified in (Wang & Zhe, 2022b), but with rank set to 10 to align with the desired 1332
- BBs. The algorithm was trained for 500 epochs across 10 folds. BBs were extracted from the two views of the " Z_{arr} " matrix during the last epoch. The first view was reweighted using the weights obtained from the second view of " Z_{arr} " for each state and BB. Temporal traces were then extracted from the " U_{arr} " matrix to calculate the trace of each BB under each state.
- For NNDTN (discrete-time NN decomposition with nonlinear dynamics, as implemented by (Wang & Zhe, 2022a)), we concatenated individual components of " v_{i_n} " over states over the number of BBs across all time points with re-weighting by " U_{vec} ". The traces were then obtained by optimizing the BBs' activity to minimize the distance between the reconstruction and the original tensor.

Post-processing steps applied to baselines' BBs and traces to align them with the ground truth:

1343 1344 • *Synthetic Data*:

1345 To assess and compare the results of these alternative methods against the ground truth BBs and traces, we initially 1346 normalized the BBs to fit the range of the ground truth BBs, applied sparsity using hard-thresholding such that the 1347 identified BBs from each method will present similar sparsity level to that of the ground truth, and then reordered the 1348 BBs to maximize the correlations of their temporal traces with the ground truth traces. This alignment was necessary 1349 since SiBBIInGS is insensitive to the ordering of BBs. For the correlation comparisons ($\rho(\mathbf{A}, \hat{\mathbf{A}})$), we examined the 1350 correlation between the BBs, as well as their temporal traces, in comparison to the ground truth. Recognizing that 1351 correlation might not be the most suitable metric for sparse BBs comparison, we further evaluated the performance 1352 using the Jaccard index as well.

13531354 • EEG and Trends Experiments:

1355 Similar to the synthetic data scenario, in the EEG and Trends experiments, we compared the identified components 1356 with outcomes generated by different tensor and matrix factorization methods. However, since these experiments are 1357 built on real-world data, in these case, no ground truth exists for the ensembles, as the 'real' ensembles are unknown 1358 and hence the evaluation of the identified structures is qualitative. Specifically, after applying the baselines to the 1359 EEG/Trends data, we extracted the BBs (A) by the following: In the cases of global and local PCA, these BBs were 1360 treated as the Principal Components (PCs). In the PARAFAC and Tucker tensor-factorization methods, they were 1361 considered the first factor (factors[0] from the tensorly output), weighted by the relevant components from the third 1362 factor (the states axis, factors[2]). We then performed the following steps: 1) Normalized the matrices to ensure that 1363 each BB had a similar absolute sum of its columns, resulting in BBs of comparable magnitudes for state comparison, 1364 and 2) Introduced artificial sparsity into the matrices through hard thresholding, aiming to achieve the same level of 1365 sparsity observed in SiBBIInGS for each state. As seen in Figure 5, in the EEG experiment, the baselines failed to detect the emergence of BBs around electrode F8, resulting in widespread non-specific clusters; As seen in Figure 10, 1366 1367 in the Trneds experiment, these methods produced less meaningful BBs than SiBBIInGS.

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Figure 13. Additional Figures for the Neural Recordings Experiment. A The identified BBs for the different states. While there is 1417 clear consistency, slight modifications can be observed across states, capturing the natural variability in neural ensembles corresponding 1418 to different tasks. B Temporal traces of the identified BBs, shown with a 90% confidence interval (background color), and all trials are 1419 plotted in light gray. The color corresponds to the state color used in Figure 4. We observe adaptation over the states as well as differences 1420 between the temporal traces of BBs within a given state. The third BB exhibits significantly lower activity compared to the others (see 1421 also Figure 4), suggesting that it might capture general background trends or noise. C Within and between temporal trace correlations 1422 (averaged over 100 bootstrapped examples) with standard error, colored according to the BB color (as used in Fig. 4), and transparency 1423 representing the strength of the between (opaque) and within (less opaque) correlations. D Example of the correlations between each pair 1424 of BBs within the 1-st state (0°) . This shows that while some BBs are orthogonal, others are not. E Example of within-state correlations 1425 between each pair of temporal traces of the BBs within the 1-st trial of the 1-st state (0°) , showing that the temporal traces are neither 1426 orthogonal nor overly correlated. F Accuracy in predicting the state using only the temporal traces of that state as input (colored by the state color). While the random accuracy would be $1/\text{length}(\text{labels}) = \frac{1}{8} = 0.125$, the achieved accuracies are significantly higher for all 1427 1428 states.