Bounded Self-Weights Estimation Method for Non-Local Means Image Denoising Using Minimax Estimators

Minh Phuong Nguyen and Se Young Chun, *Member, IEEE*

*Abstract***— A non-local means (NLM) filter is a weighted average of a large number of non-local pixels with various image intensity values. The NLM filters have been shown to have powerful denoising performance, excellent detail preservation by averaging many noisy pixels, and using appropriate values for the weights, respectively. The NLM weights between two different pixels are determined based on the similarities between two patches that surround these pixels and a smoothing parameter. Another important factor that influences the denoising performance is the self-weight values for the same pixel. The recently introduced local James-Stein type center pixel weight estimation method (LJS) outperforms other existing methods when determining the contribution of the center pixels in the NLM filter. However, the LJS method may result in excessively large self-weight estimates since no upper bound is assumed, and the method uses a relatively large local area for estimating the self-weights, which may lead to a strong bias. In this paper, we investigated these issues in the LJS method, and then propose a novel local self-weight estimation methods using direct bounds (LMM-DB) and reparametrization (LMM-RP) based on the Baranchik's minimax estimator. Both the LMM-DB and LMM-RP methods were evaluated using a wide range of natural images and a clinical MRI image together with the various levels of additive Gaussian noise. Our proposed parameter selection methods yielded an improved bias-variance trade-off, a higher peak signal-to-noise (PSNR) ratio, and fewer visual artifacts when compared with the results of the classical NLM and LJS methods. Our proposed methods also provide a heuristic way to select a suitable global smoothing parameters that can yield PSNR values that are close to the optimal values.**

*Index Terms***— James-Stein estimator, minimax estimator, non-local means, center pixel weight, bounded self-weight, image denoising.**

I. INTRODUCTION

IMAGE denoising is a fundamental task in image processing, low-level computer vision, and medical imaging algoing, low-level computer vision, and medical imaging algorithms. The goal of denoising is to suppress image noise

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Minh Phuong Nguyen was with the School of Electrical and Computer Engineering (ECE), Ulsan National Institute of Science and Technology (UNIST), Ulsan, 44919, South Korea. Now he is with Section of Radiation, Detection and Medical Imaging, Delft University of Technology, Mekelweg 15, 2629 JB Delft, The Netherlands (e-mail: m.p.nguyen@tudelft.nl).

Se Young Chun is with the School of ECE, UNIST, Ulsan, 44919, South Korea (e-mail: sychun@unist.ac.kr).

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when restoring desired details using prior information about the images. For example, based on prior information regarding "smooth images", a simple filter, such as a Gaussian filter, can be designed as a weighted average of the image intensities of the pixels in the local neighborhood with non-adaptive weights. However, this type of filter blurs the edges and details of images because these features are not captured in the assumed prior information. Many edge-preserving denoising methods have been proposed, including bilaterial filters [1], [2], anisotropic diffusion [3], non-local means (NLM) filters [4], [5], collaborative filters (BM3D) [6], and total variation filters [7]. Many filters, including bilaterial filters, anisotropic diffusion, and NLM filters (but, not BM3D, see [8]), can be represented as the weighted averages of adaptive weights or adaptive smoothing [9]. It should be noted that it is important to select appropriate weights in these types of filters in order to obtain improved denoised image quality [8].

Classical NLM filters use the similarities between two local patches in a noisy image to determine the weights in nonlocal adaptive smoothing [4]. The NLM weights are obtained by first calculating the Euclidean distance between the two local patches, which is denoted *d*, and then by evaluating $\exp(-d^2/h^2)$, where *h* is a smoothing parameter. This method allows higher weights to be assigned to pixels with similar patches so that edges and details can be preserved through non-local weighted averaging.

There are four different factors that determine the output image quality of a NLM filter in terms of weights. 1) The first factor is the similarity measure *d*. The Euclidean distance is a usual choice [4], but other similarity measures have also been proposed, such as hypothesis testing with adaptive neighborhoods [10], principal component analysis (or the subspace based method) [11], [12], blockwise aggregation [13], rotationinvariant measures [14]–[16], shape-adaptive patches [17], and patch-based similarities with adaptive neighborhoods [18]. In multimodal medical imaging, inaccurate weights for noisy molecular images were enhanced by using additional high quality anatomical images [19], [20]. 2) The second factor is the strategy for determining the smoothing parameter *h*. Optimization strategies have been developed based on Stein's unbiased risk estimation (SURE) method for NLM with Gaussian noise [21], [22], NLM with Poisson noise [23], and blockwise NLM with Gaussian noise [24]. 3) The third factor is in selecting the function to use to determine the weights, such as $exp(-x^2)$. Other functions have also been proposed to calculate the weights, such as compact support func-

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tions [25], [26] and statistical distance functions [27], [28]. 4) The last factor, which is the focus of this article, is in how best to determine the self-weights for the same pixel in the input and output images.

The NLM weights for two different pixels are essentially determined by the distance between the two noisy local patches around these pixels. However, the weights for the same pixel, or the self-weights, are not affected by the noise in the patches and the distance is always 0. For an extremely noisy image, the self-weights will be relatively too large when compared to the other weights, which will cause the filter output to be almost the same as the input noisy image. Therefore, the use of appropriate self-weight values can significantly affect the quality of the denoised image. Many researchers have investigated strategies for determining the self-weights, which are also known as center pixel weights, in order to alleviate the so-called "rare patch effect." For the classical NLM filter proposed by Buades *et al.*, the selfweights were set to be either one or the maximum weight in a neighborhood [4]. This strategy guaranteed that at least one or two of the largest weights would be the same. Doré and Cheriet also used the maximum weight in a neighborhood as the self-weight, but only if that maximum weight was large [29]. Brox and Cremers proposed a method to have at least *n* number of the weights to be the same [30], and Zimmer *et al.* considered the self-weight to be a free parameter during the estimation process [31]. Salmon developed a SURE-based method for determining the self-weights that accounted for the noise [32].

Recently, Wu *et al.* proposed a method to determine the self-weights using a James-Stein (JS) type estimator [33]. The idea of that work was to use a JS estimator to determine the reparametrized self-weight in a local neighborhood (called the local JS estimator (LJS)). The LJS method yielded the best peak signal-to-noise ratio (PSNR) results when compared to other existing self-weight selection methods [4], [32]. However, the method had some limitations. First, the LJS could yield self-weights that were theoretically much larger than 1 because no upper bound for the self-weights was assumed, and this may lead to severe rare patch artifacts. The JS estimator does not guarantee its optimality for bounded shrinkage parameters. Second, the original LJS method was tested with a relatively large local neighborhood when determining a self-weight because it was assumed that the selfweights were the same in the local neighborhood. However, the problem is that the selection of a local neighborhood size that is too large may introduce a strong bias into the resulting denoised images.

In this article, we investigate the original LJS method in terms of the local neighborhood size for self-weight estimation and the potential for excessive self-weight estimation when no upper bound is applied on the self-weight. We then propose novel self-weight estimation methods for NLM that account for bounded self-weights using Baranchik's minimax estimator [34], called local minimax self-weight estimation with direct bound (LMM-DB) and with reparametrization (LMM-RP). We evaluated our proposed methods using performance criteria including PSNR, the bias-variance tradeoff curve and visual quality assessment with a wide range of natural images and a real patient MRI image with various noise levels. We compared the performance of our proposed methods with a classical NLM filter using self-weights of 1 [4] and the state-of-the-art LJS method, which has already been shown to be the best among all other previous self-weight determination methods [33].

This article is an extension of a work that was presented at the 2016 IEEE International Symposium on Biomedical Imaging (ISBI) [35], and goes into more depth regarding the theory of the minimax estimator and provides an evaluation of the methods using a significantly larger image dataset.

This paper is organized as follows. Section II reviews the classical NLM filter and revisits the LJS method. Section III investigates the LJS method in terms of the local neighborhood size for self-weight estimation and the potential for excessively large self-weight estimates. Then, Section IV proposes novel LMM-DB and LMM-RP methods using Baranchik's minimax estimator in order to overcome two limitations of the LJS method. Section V illustrates the performance of our proposed methods by providing our simulation results. Lastly, Sections VI and VII discuss and then conclude this paper, respectively.

II. REVIEW OF THE LOCAL JAMES-STEIN SELF-WEIGHT ESTIMATION METHOD FOR THE NLM FILTER

In this section, we will briefly review both the classical NLM method proposed by Buades *et al.* [4] and the LJS selfweight selection method proposed by Wu *et al.* [33].

A. Reviewing the Classical Non-Local Means Filter

Let us assume that an image **x** is contaminated by noise **n**, which produces a noisy image **y**:

$$
y = x + n \tag{1}
$$

where **n** is zero-mean white Gaussian noise with standard deviation σ . The NLM filtered value at the pixel *i* is the weighted average of all pixels in a search region Ω_i :

$$
\hat{x}_i = \frac{\sum_{j \in \Omega_i} w_{i,j} y_j}{\sum_{j \in \Omega_i} w_{i,j}}
$$
\n(2)

where y_i is the *i*th element of **y**, $w_{i,j}$ is the weight between the *i*th and *j*th pixels, and Ω_i is the set of all pixels in an area around the *i*th pixel, which could be an entire image. The similarity weight of the classical NLM is defined as:

$$
w_{i,j} = \exp\left(\frac{-\left\|\mathbf{P}_i\mathbf{y} - \mathbf{P}_j\mathbf{y}\right\|^2}{2\left\|\mathbf{P}\right\|h^2}\right) \tag{3}
$$

where P_i is an operator used to extract a square-shaped patch centered at the *i*th pixel, $\|\cdot\|$ is an l_2 norm, $|\mathbf{P}|$ is the number of pixels within a patch, and *h* is a global smoothing parameter. Equation (3) implies that the self-weights $w_{i,i}$ are always equal to 1. Previous works on self-weights have shown that good strategies for determining the self-weights also affect the image quality of the NLM filtering [4], [29], [32], [33].

B. Reviewing Local James-Stein Self-Weight Estimation

The LJS method was proposed in order to determine $w_{i,i}$ as follows [33]. First, (2) was decomposed into two terms:

$$
\hat{x}_i = \frac{W_i}{W_i + w_{i,i}} \hat{z}_i + \frac{w_{i,i}}{W_i + w_{i,i}} y_i
$$
\n(4)

where $W_i = \sum_{j \in \Omega_i \setminus \{i\}} w_{i,j}$ and

$$
\hat{z}_i = \sum_{j \in \Omega_i \setminus \{i\}} w_{i,j} y_j / W_i. \tag{5}
$$

The terms \hat{z}_i do not contain $w_{i,i}$. Then, the LJS method reparametrized (4) using

$$
p_i = \frac{w_{i,i}}{W_i + w_{i,i}}\tag{6}
$$

so that (4) became:

$$
\hat{x}_i = (1 - p_i)\,\hat{z}_i + p_i\,y_i.\tag{7}
$$

The problem of estimating the self-weights $w_{i,i}$ became the problem of estimating p_i . Lastly, the JS estimator [36], [37] for *pi* was proposed:

$$
p_i^{\text{LJS}} = 1 - \frac{(|\mathbf{B}| - 2)\sigma^2}{\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|^2}
$$
(8)

where \mathbf{B}_i is an operator used to extract a square-shaped neighborhood centered at the *i*th pixel, |**B**| is the number of pixels within that neighborhood, and σ is the known noise level.

Equation (8) implies that $p_i^{\text{LJS}} \in (-\infty, 1]$. Since the weights are non-negative, it was proposed to use the zero-lower bound for p_i^{LJS} as follows [33]:

$$
\hat{x}_i^{\text{LJS}_+} = \left(1 - p_i^{\text{LJS}_+}\right)\hat{z}_i + p_i^{\text{LJS}_+}y_i \tag{9}
$$

where

$$
p_i^{\text{LJS}_+} := [p_i^{\text{LJS}}]_+ = \left[1 - \frac{(|\mathbf{B}| - 2)\sigma^2}{\left\|\mathbf{B}_i\mathbf{y} - \mathbf{B}_i\hat{\mathbf{z}}\right\|^2}\right]_+\tag{10}
$$

and $[s]_+ := \max(s, 0)$. Wu *et al.* also mentioned that a userdefined upper bound for *pi* can be used, but did not investigate further [33]. It should be noted that the JS estimator does not guarantee its optimality when bounding p_i^{LJS} in (8).

III. LIMITATIONS OF THE LOCAL JAMES-STEIN SELF-WEIGHT ESTIMATION FOR THE NLM FILTER

We now investigate two limitations of the original LJS method [33] in terms of the size of local neighborhoods for self-weight estimation, and the potential for excessive selfweight estimation.

A. Size of Local Neighborhood for Self-Weight Estimation

In the method described in [33], there are two implicit steps required in order to obtain the LJS self-weight estimator (10). The first step is to choose a local set of pixels around the *i*th pixel, referred to as set Ω_i^B , that correspond to the operator \mathbf{B}_i , and assume that:

$$
\hat{x}_j = (1 - p_i)\,\hat{z}_j + p_i\,y_j, \quad j \in \Omega_i^B. \tag{11}
$$

Fig. 1. Bias-variance curves (cameraman example) for the classical NLM and LJS methods $(LJS₊)$ for different sizes of local neighborhoods (B). The curves were plotted while varying the smoothing parameter h ($log_2 h \in [1.8, 3.2]$).

Based on the works of Stein [36] and James and Stein [37], if $|\mathbf{B}| \geq 3$, then for a neighborhood Ω_i^B extracted using \mathbf{B}_i ,

$$
\hat{x}_j = \left(1 - p_i^{\text{LJS}_+}\right)\hat{z}_j + p_i^{\text{LJS}_+}y_j, \quad j \in \Omega_i^B. \tag{12}
$$

is a dominant estimator for x_j "locally" in Ω_i^B . The LJS method used the zero lower bound when estimating p_i in order to obtain a realistic non-negative self-weight value. This was also a good choice in terms of the estimator performance since the positive part of the JS estimator is dominant over the original JS estimator, according to the works of Baranchik [34], [38] and Efron and Morris [39].

The second implicit step is to assign the resulting p_i^{LJS} to p_i in (7) for only the single pixel *i* so that:

$$
\hat{x}_i^{\text{LJS}} = \left(1 - p_i^{\text{LJS}}\right)\hat{z}_i + p_i^{\text{LJS}}y_i. \tag{13}
$$

Wu *et al.* evaluated the LJS method with $|\mathbf{B}| = 15 \times 15$ [33], which seems relatively large.

Based on this implicit two-step interpretation, we can surmise that using a smaller size of |**B**| may be more desirable for obtaining a less biased estimate of p_i since the assumption of having the same p_i in Ω_i^B is less likely to be true for larger sizes of Ω_i^B . Figure 1 confirms our conjecture. The bias-variance curves of the LJS method yielded better biasvariance trade-offs than those in the classical NLM method for both large local neighborhoods with a half window size $B = 7$ $(|\mathbf{B}| = 15 \times 15)$ and small local neighborhoods with $B = 2$ $(|\mathbf{B}| = 5 \times 5)$. However, using larger local neighborhood sizes for estimating *pi* yielded stronger biases than those estimated using smaller sizes for the same level of variance.

B. Excessively Large Self-Weight Estimates

In the LJS method for determining the self-weights by estimating values for p_i [33], it is theoretically possible that the self-weights have excessively high values. For example, (6) suggests that if $p_i = 1$ and $W_i > 0$, then $w_{i,i} \gg 1$. Slight artifacts were observed in [33] in the background area that were potentially caused by excessive self-weight estimates when a relatively larger neighborhood size $|\mathbf{B}| = 15 \times 15$ was used. We observed a significantly higher degree of degradation in the visual image quality in the background area when the

Fig. 2. Denoised image of the cameraman example using the original LJS method [33] with no upper bound for the self-weights (top left), estimated *pi* values (top right), calculated W_i (bottom left), and resulting self-weights $(w_{i,i})$ showing excessive self-weights (bottom right). B = 2 and $\sigma = 10$.

size of $|\mathbf{B}|$ in (8) was small, as shown in the image in the top left figure of Fig. 2.

We investigated this issue using an example of the cameraman image that was denoised using the LJS method [33], but with a smaller neighborhood size $|\mathbf{B}| = 5 \times 5$. For areas with more details, such as edges and textures, large *pi* values were estimated and yielded large self-weights, as shown in the top right figure of Fig. 2. However, since the values for W_i were also very small in these areas, as shown in the bottom left image in Fig. 2, the resulting self-weight map yielded values close to 1 in the areas with details as shown in the bottom right image in Fig. 2.

In contrast, for areas with almost no details, such as those with a flat intensity background, relatively smaller p_i values were estimated, some of which were much larger than 0 while the rest were closer to 0, as shown in the top right image in Fig. 2. However, since the *Wi* values for the flat areas were relatively large, as shown in the bottom left image in Fig. 2, some of the estimated *pi* values obtained using the LJS method (LJS+) were estimated to yield excessively large self-weights that were much larger than 1, as shown in the bottom right image of Fig. 2. Consequently, these excessively large self-weights caused severe rare patch artifacts in the filtered image, which resulted in visual quality degradation, as observed in the top left image of Fig. 2.

IV. LOCAL MINIMAX ESTIMATION METHODS FOR UPPER BOUNDED SELF-WEIGHTS IN A NLM FILTER

In this section, we propose two local upper bounded selfweight estimation methods that use Baranchik's minimax estimator [34].

A. Bounded Self-Weights

It is usually assumed that the self-weights satisfy $w_{i,i}$ ∈ [0, 1]. However, there are many possible upper bounds for the self-weights, including 1 [4] or some positive value that is possibly less than 1 based on SURE [32]. In this article, two different upper bound values $w_{i,i}^{\text{max}}$ for the self-weights were evaluated such that $0 \leq w_{i,i} \leq w_{i,i}^{\max}$. One upper bound was:

$$
w_{i,i}^{\max-one} = 1,\tag{14}
$$

which is the usual choice for the self-weights in the classical NLM method [4]. The other upper bound was:

$$
w_{i,i}^{\max - \text{stein}} = \exp\left(-\sigma^2/h^2\right),\tag{15}
$$

which was motivated by the SURE-based NLM self-weights [32]. We assume that σ is known and *h* is pre-determined, which means that the upper bound for the self-weights can also be determined in advance. Equation (15) takes the noise level into account. As σ is smaller, the maximum self-weight in (15) is closer to one. It should be noted that the difference between (15) and (14) will be greater at higher noise levels.

Since p_i is estimated instead of $w_{i,i}$, it is necessary to derive the range of p_i that corresponds to $0 \le w_{i,i} \le w_{i,i}^{\max}$. From (6), the derivative of p_i with respect to $w_{i,i}$ is nonnegative as follows:

$$
\frac{d}{dw_{i,i}}p_i = \frac{W_i}{(W_i + w_{i,i})^2} \ge 0
$$

since $W_i \geq 0$. Therefore, p_i is a non-decreasing function of $w_{i,i}$ and for $0 \le w_{i,i} \le w_{i,i}^{\max}$, the range of p_i will be

$$
0 \le p_i \le \frac{w_{i,i}^{\max}}{W_i + w_{i,i}^{\max}} =: p_i^{\max} \le 1.
$$

Note that if $W_i = 0$, then $p_i^{\text{max}} = 1$. The estimator $p_i^{\text{LJS}_+}$ in (10) automatically guarantees that $0 \leq p_i \leq 1$ if $|\mathbf{B}| \geq 2$. However, since $W_i > 0$ generally holds for most real images with noise, it is necessary to constrain p_i to be less than or equal to the upper bound p_i^{max} , which is usually less than one.

B. Local Minimax Self-Weight Estimation With Direct Bound

Enforcing the upper limit p_i^{max} on the estimated p_i in (10) using $\min(p_i^{\text{LJS}_+}, p_i^{\text{max}})$ breaks the optimality of the JS estimator if p_i^{\max} < 1. In this article, we propose using Baranchik's minimax estimator [34] to incorporate bounded self-weights into the estimator (see Baranchik [34], Erfon and Morris [39], and Strawderman [40] for more details on this minimax estimator).

Theorem 1 (Baranchik): For $y \sim \mathcal{N}_r(x, \sigma^2 I)$, $r \geq 3$, and $L(\mathbf{x}, \hat{\mathbf{x}}) = ||\mathbf{x} - \hat{\mathbf{x}}||$, an estimator of the form $\hat{\mathbf{x}} = q\mathbf{y}$ where

$$
q = \left[1 - c\left(\|\mathbf{y}\|\right) \frac{\sigma^2 (r-2)}{\|\mathbf{y}\|^2}\right] \tag{16}
$$

is the minimax, provided that:

(i) $0 \le c$ ($||y||$) ≤ 2 and

(ii) the function $c(\cdot)$ is nondecreasing.

Here **y** shrinks toward **0** which is the initial estimate of **x**.

Fig. 3. Graphical illustrations of the original and positive part JS estimators without upper bounds, and the proposed minimax self-weight estimators with upper bounds in terms of c ($||\mathbf{s}||$) vs. $||\mathbf{s}||$. (a) Original and positive-part JS estimators. (b) Proposed minimax estimators with bounds.

The original JS estimator and its positive part are special cases of Baranchik's minimax estimator. For the original JS estimator (8):

$$
c\left(\left\|\mathbf{s}\right\|\right) = 1,\tag{17}
$$

where $\mathbf{s} = \mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}$ so that both conditions (i, ii) of the Baranchik's theorem are satisfied. In the positive part estimator (9), it can be shown that:

$$
c\left(\left\|\mathbf{s}\right\|\right) = \begin{cases} \frac{\left\|\mathbf{s}\right\|^2}{\sigma^2(r-2)}, & 0 \le \left\|\mathbf{s}\right\| \le \mathrm{Y}_1\\ 1, & \text{otherwise} \end{cases} \tag{18}
$$

where $Y_1 := \sigma \sqrt{r-2}$. The original and positive part JS estimators are illustrated in Fig. 3 (a).

We propose a new local minimax self-weight estimation method that uses a direct bound with a specific upper-bound value, as follows:

$$
p_i^{\text{LMM}-\text{DB}} := \min(p_i^{\text{LJS}_+}, p_i^{\text{max}}). \tag{19}
$$

This estimator is minimax under certain conditions that can be derived using Baranchik's minimax estimator theorem. According to this theorem, this operation can be interpreted as follows:

$$
c(||\mathbf{s}||) = \begin{cases} \frac{||\mathbf{s}||^2}{\sigma^2(r-2)}, & 0 \le ||\mathbf{s}|| \le Y_1\\ 1, & Y_1 < ||\mathbf{s}|| \le Y_2\\ \frac{||\mathbf{s}||^2 (1 - p^{\max})}{\sigma^2(r-2)}, & Y_2 < ||\mathbf{s}|| \end{cases}
$$
 (20)

where $Y_2 := \sigma \sqrt{(r-2)/(1-p^{\text{max}})}$. We call this a local minimax self-weight estimator using direct bound (LMM-DB), which is illustrated in Fig. 3 (b) bound (LMM-DB), which is mus
where Y₄ := $σ\sqrt{2(r-2)/(1-p^{max})}$.

However, note that LMM-DB is not minimax for $||s|| > Y_4$. Fortunately, $\|\mathbf{s}\| = \|\mathbf{B}_i\mathbf{y} - \mathbf{B}_i\hat{\mathbf{z}}\|$ can be limited by adjusting the smoothing parameter *h* by making it smaller so that all $\|\mathbf{s}\|$ \leq Y₄ and c ($\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|$) \leq 2. Then, the LMM-DB becomes "practically" a minimax estimator. Let us denote the maximum *h* that satisfies $\|\mathbf{s}\| \leq Y_4$ as h^{max} .

In this case, a question can be raised: will the optimal value for *h* fall into the range of *h* that satisfies $\|\mathbf{s}\| \leq Y_4$? Interestingly, our simulations with many natural images showed that the optimal smoothing parameter *h*[∗] based on the true images is very close to h^{max} . This is because the LMM-DB yielded $p^{\text{max}} \to 1$ so that $Y_2 \to \infty$, and almost all $\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|$ were less than or equal to Y_4 . Therefore, $p_i^{\text{LMM}-\text{DB}}$ is "practically" a minimax value based on Baranchik's theorem for many natural images. Moreover, the LMM-DB method may provide a way to choose the optimal global smoothing parameter value *h* without knowing the underlying true image. We empirically investigate this issue in Section V.

C. Local Minimax Self-Weight Estimation With Re-Parametrization

The LMM-DB algorithm set p to be the same p^{\max} for a wide range of $\|\mathbf{B}_i\mathbf{y} - \mathbf{B}_i\hat{\mathbf{z}}\|$ values. We now propose another new method, called the local minimax self-weight estimation with reparametrization (LMM-RP) method, that assigns different *p* values for different $\|\mathbf{B}_i\mathbf{y} - \mathbf{B}_i\hat{\mathbf{z}}\|$.

We reparametrized p_i in (7) in the following way:

$$
\begin{aligned}\n\hat{x}_i &= \hat{z}_i (p_i / p_i^{\max}) p_i^{\max} (y_i - \hat{z}_i) \\
&= \hat{z}_i + p_i^{\mathrm{T}} (y_i^{\mathrm{T}} - \hat{z}_i^{\mathrm{T}})\n\end{aligned} \tag{21}
$$

$$
= (1 - p_i^{\max})\hat{z}_i + \hat{z}_i^{\text{T}} + p_i^{\text{T}}(y_i^{\text{T}} - \hat{z}_i^{\text{T}})
$$
 (22)

where $\hat{z}_i^{\mathrm{T}} = p_i^{\max} \hat{z}_i$, $y_i^{\mathrm{T}} = p_i^{\max} y_i$, and

$$
p_i^{\rm T} = \frac{1}{p_i^{\max}} \frac{w_{i,i}}{W_i + w_{i,i}}.
$$
 (23)

Note that for $0 \le w_{i,i} \le w_{i,i}^{\max}$, p_i^T is an increasing function of $w_{i,i}$ and the range of p_i^T is $0 \leq p_i^T \leq 1$. We propose to use the positive part of the JS estimator to estimate the reparametrized p_i^T , as follows:

$$
p_i^{\text{T,LS}_{+}} = \left[1 - \frac{(|\mathbf{B}| - 2) (p_i^{\text{max}})^2 \sigma^2}{\|\mathbf{B}_i \mathbf{y}^{\text{T}} - \mathbf{B}_i \hat{\mathbf{z}}^{\text{T}}\|^2}\right]_+ = \left[1 - \frac{(|\mathbf{B}| - 2) \sigma^2}{\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|^2}\right]_+ = p_i^{\text{LS}_{+}}.\tag{24}
$$

This method is equivalent to using a multiplicative factor p_i^{\max} for the original JS shrinkage (9):

$$
\hat{x}_i^{\text{LMM-RP}} = (1 - p_i^{\text{LMM-RP}})\hat{z}_i + p_i^{\text{LMM-RP}}y_i \qquad (25)
$$

where

$$
p_i^{\text{LMM-RP}} = p_i^{\max} \left[1 - \frac{(|\mathbf{B}| - 2) \sigma^2}{\left\| \mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}} \right\|^2} \right]_+ . \tag{26}
$$

This proposed LMM-RP estimator is not dominant when estimating x_i , but rather is dominant when estimating $p_i^{\max} x_i$, as shown in (22). Thus, the positive part JS estimator does not guarantee that the LMM-RP is dominant.

Baranchik's minimax estimation theorem can be used to analyze the LMM-RP estimator as follows:

$$
c\left(\|\mathbf{s}\|\right) = \begin{cases} \frac{\|\mathbf{s}\|^2}{\sigma^2(r-2)}, & 0 \le \|\mathbf{s}\| \le Y_1\\ \frac{\|\mathbf{s}\|^2(1-p^{\max})}{\sigma^2(r-2)} + p^{\max}, & Y_1 < \|\mathbf{s}\| \end{cases}
$$
(27)

where if $\|s\|$ is $Y_3 := \sigma \sqrt{(2-p^{\max})(r-2)/(1-p^{\max})}$, then c ($\|\mathbf{s}\|$) = 2. The LMM-RP method is also illustrated in Fig. 3 (b), and is minimax if $\|\mathbf{s}\| \leq Y_3$. The global smoothing parameter *h* can be adjusted so that this condition is satisfied for different images. As in the case of the LMM-DB, it turns out that the optimal global smoothing parameter *h*[∗] and the upper bound *h* that satisfies $\|\mathbf{s}\| \leq Y_3$ are also very close to each other when the LMM-RP method is applied to many natural images. Therefore, the LMM-RP method is "practically" a minimax. The following table summarizes the LJS self-weight estimation method and our proposed LMM-based self-weight estimation methods.

Summary of Self-Weight Estimation Methods

LJS₊ [33]:
\n
$$
p_i^{\text{LJS}_+} = [1 - (|\mathbf{B}| - 2) \sigma^2 / ||\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}||^2]_+
$$
\n
$$
\hat{x}_i^{\text{LJS}_+} = (1 - p_i^{\text{LJS}_+})\hat{z}_i + p_i^{\text{LJS}_+} y_i
$$
\nLMM – DB:
\n
$$
p_i^{\text{LMM}-\text{DB}} = \min(p_i^{\text{LJS}_+}, p_i^{\text{max}})
$$
\n
$$
\hat{x}_i^{\text{LMM}-\text{DB}} = (1 - p_i^{\text{LMM}-\text{DB}})\hat{z}_i + p_i^{\text{LMM}-\text{DB}} y_i
$$
\nLMM – RP:
\n
$$
p_i^{\text{LMM}-\text{RP}} = p_i^{\text{LJS}_+} p_i^{\text{max}}
$$
\n
$$
\hat{x}_i^{\text{LMM}-\text{RP}} = (1 - p_i^{\text{LMM}-\text{RP}})\hat{z}_i + p_i^{\text{LMM}-\text{RP}} y_i
$$

V. SIMULATION RESULTS

A. Simulation Setup

Ten natural images¹ (cameraman, lena, montage, house, pepper, barbara, boat, hill, couple, fingerprint) and five images from the SUN database² (abbey, airplane cabin, airport terminal, alley, amphitheater) were used in our study as noisefree images (128 \times 128, 256 \times 256, or 512 \times 512 pixels, 8 bits). A real patient MRI (512×512 pixels, 8 bits) that was acquired and processed under institutional review board (IRB) approved protocols was also used. White Gaussian noise was added to each input image with various standard deviations $\sigma \in \{10, 20, 40, 60\}.$

All algorithms were implemented using MATLAB R2015b (The Mathworks, Inc., Natick, MA, USA). The patch size and search window size of the NLM filter were chosen to be 7×7 and 31×31 , respectively, which were the same as those used in [33]. Both the state-of-the-art LJS algorithm and the proposed algorithms were tested using $B = 1, \dots, 9$ where $|\mathbf{B}| = (2B + 1)^2 > 3$.

The global smoothing parameter *h* was chosen empirically to yield the best PSNR:

PSNR
$$
(\hat{\mathbf{x}})
$$
 = 10 log₁₀ $\frac{255^2}{\|\hat{\mathbf{x}} - \mathbf{x}\|^2 / N}$, (28)

where *N* is the size of the image. In addition to the PSNR, the mean bias vs. the mean variance trade-off curves were used as performance measures for the different smoothing parameter values *h*:

$$
\overline{\text{bias}^2} = \frac{1}{N} \sum_{i=1}^{N} (\bar{x}_i - x_i)^2,
$$
 (29)

$$
\overline{\text{var}} = \frac{1}{N} \sum_{i=1}^{N} \frac{1}{k-1} \sum_{j=1}^{k} (\hat{x}_{ij} - \bar{x}_i)^2, \quad (30)
$$

where k is the number of realizations $(20 \text{ in our simulation})$, \hat{x}_{ij} is the *j*th estimation at the *i*th pixel, and \bar{x}_i is the mean of \hat{x}_{ij} , as given by:

$$
\bar{x}_i = \frac{1}{k} \sum_{j=1}^k \hat{x}_{ij}.
$$

A visual quality assessment was also performed.

B. Performance Studies Using the PSNR

In order to estimate values of *pi* for a fixed neighborhood size *B*, the optimal NLM smoothing parameter h^* was determined such that the PSNR was maximized. In our proposed methods, the two maximum self-weights in (14) and (15) were used. The LMM-DB and LMM-RP methods given by (14) are denoted $LMM - DB^{one}$ and $LMM - RP^{one}$, while the LMM-DB and LMM-RP methods given by (15) are denoted $LMM - DB^{stein}$ and $LMM - RP^{stein}$. Table I summarizes the quantitative PSNR results for the 16 images with 4 different noise levels. When $B = 7$, our proposed LMM-DB and LMM-RP methods based on Baranchik's minimax estimator yielded much better PSNR results than did setting the self-weight to one in the classical NLM method [4], and comparable PSNR values to the LJS method based on the JS estimator [33]. When $B = 2$, our proposed LMM-DB and LMM-RP methods yielded better PSNR values than did the LJS.

For the five examples of lena, house, peppers, barbara, boat with $\sigma = 20$, PSNRs of our proposed methods (global smoothing parameter and fixed neighborhoods, but adaptive self-weight) were $0.72 \sim 0.97$ dB better than classical NLM. In [10], it is reported that for the same five examples with the same level of noise, the work of Kervrann et al. (fixed self-weight, but local smoothing parameters and adaptive neighborhoods) yielded $0.99 \sim 1.55$ dB better PSNR than classical NLM. Self-weights, local smoothing parameter, neighborhoods size are important factors in the NLM filter to determine output image quality.

¹Available online at: http://www.cs.tut.fi/~foi/GCF-BM3D/BM3D_images. zip as the date of 16 Nov. 2015. 2Available online at: http://vision.princeton.edu/projects/2010/SUN/ as the

date of 16 Sep. 2016.

TABLE I PSNR (dB) SUMMARY (MEAN ± STANDARD DEVIATION) FOR VARIOUS NATURAL IMAGES

		$B = 2$			$B = 7$							
	Q.	Classical NLM	LJS ₊	$LMM-DB$ ^{one}	$LMM-RPone$	LMM-DB ^{stein} LMM-RP ^{stein}		LJS ₊	$LMM-DB$ ^{one}	LMM-RP one	$LMM-DB$ ^{stein}	LMM-RP stein
cameraman	10	32.42 ± 0.034		33.12 ± 0.031 33.32 ± 0.030 33.29 ± 0.029 33.17 ± 0.029 33.04 ± 0.030						32.98 ± 0.035 33.10 ± 0.036 33.05 ± 0.037 32.98 ± 0.038 32.85 ± 0.040		
	20	28.48 ± 0.052		29.12 ± 0.052 29.46 \pm 0.056 29.29 \pm 0.056 29.27 \pm 0.060 28.97 \pm 0.052						29.32 ± 0.062 29.34 \pm 0.059 29.04 \pm 0.054 29.11 \pm 0.058 28.80 \pm 0.049		
	40 60	25.35 ± 0.059 23.19 ± 0.065		25.39 ± 0.083 25.89 ± 0.075 26.11 \pm 0.082 26.08 ± 0.087 25.74 ± 0.079 22.88 ± 0.062 23.39 ± 0.055 23.74 ± 0.065 23.67 ± 0.069 23.54 ± 0.061						25.98 ± 0.073 25.98 ± 0.073 25.94 ± 0.080 25.96 ± 0.083 25.63 ± 0.076 23.68 ± 0.068 23.69 ± 0.068 23.62 ± 0.071 23.63 ± 0.070 23.54 ± 0.059		
lena	10	33.90 ± 0.018		34.52 ± 0.017 34.74 ± 0.017 34.81 ± 0.018 34.77 ± 0.017 34.69 ± 0.017						34.82 ± 0.020 34.83 ± 0.020 34.80 ± 0.019 34.76 ± 0.018 34.63 ± 0.018		
	20	30.78 ± 0.031		30.90 ± 0.023 31.25 ± 0.033 31.50 ± 0.032 31.51 ± 0.032 31.31 ± 0.029						31.51 ± 0.031 31.51 ± 0.031 31.45 ± 0.028 31.48 ± 0.031 31.27 ± 0.029		
	40	27.64 ± 0.032		26.94 ± 0.028 27.74 \pm 0.033 28.08 \pm 0.029 28.08 \pm 0.029 28.06 \pm 0.028						28.10 ± 0.028 28.10 ± 0.028 28.07 ± 0.028 28.08 ± 0.029 28.06 ± 0.028		
montage	60 10	25.60 ± 0.052 34.68 ± 0.045		24.38 ± 0.040 25.66 ± 0.051 26.02 ± 0.052 26.01 ± 0.053 26.02 ± 0.054 35.19 ± 0.043 35.60 ± 0.042 35.67 ± 0.039 35.65 ± 0.042 35.55 ± 0.045						26.01 ± 0.051 26.01 ± 0.051 26.02 ± 0.054 26.02 ± 0.053 26.02 ± 0.054 35.12 ± 0.049 35.39 ± 0.046 35.38 ± 0.042 35.46 ± 0.045 35.34 ± 0.047		
	20	30.35 ± 0.088		30.74 ± 0.062 31.29 ± 0.067 31.38 ± 0.070 31.40 ± 0.078 31.06 ± 0.068						31.00 ± 0.076 31.13 ± 0.073 31.07 ± 0.068 31.18 ± 0.068 30.81 ± 0.063		
	40	26.24 ± 0.063		26.30 ± 0.072 26.98 ± 0.070 27.29 ± 0.063 27.30 ± 0.064 27.20 ± 0.061						27.01 ± 0.052 27.03 ± 0.052 27.04 ± 0.055 27.08 ± 0.053 27.00 ± 0.053		
	60	23.76 ± 0.104		23.48 ± 0.116 24.16 ± 0.113 24.61 ± 0.115 24.60 ± 0.106 24.38 ± 0.115						24.33 ± 0.092 24.33 ± 0.092 24.36 ± 0.090 24.37 ± 0.088 24.16 ± 0.095		
house	10	34.57 ± 0.038 $20\quad 31.43 \pm 0.063$		35.02 ± 0.039 35.36 ± 0.041 35.38 ± 0.039 35.34 ± 0.043 35.29 ± 0.046						35.31 ± 0.042 35.32 ± 0.043 35.25 ± 0.044 35.21 ± 0.047 35.12 ± 0.045		
	40	27.62 ± 0.044		31.54 ± 0.048 32.13 ± 0.050 32.39 ± 0.048 32.39 ± 0.050 32.19 ± 0.067 27.18 ± 0.038 27.84 ± 0.049 28.37 ± 0.037 28.37 ± 0.039 28.33 ± 0.045						32.30 ± 0.052 32.31 ± 0.054 32.26 ± 0.058 32.30 ± 0.056 32.10 ± 0.068 28.35 ± 0.041 28.35 ± 0.041 28.34 ± 0.045 28.35 ± 0.042 28.33 ± 0.045		
	60	25.01 ± 0.092		24.24 ± 0.087 25.17 ± 0.098 25.65 ± 0.095 25.65 ± 0.087 25.65 ± 0.088						25.65 ± 0.088 25.65 ± 0.088 25.65 ± 0.087 25.66 ± 0.087 25.65 ± 0.088		
peppers	10	32.62 ± 0.056		33.37 ± 0.042 33.53 ± 0.048 33.56 ± 0.049 33.49 ± 0.051 33.39 ± 0.050						33.28 ± 0.043 33.37 ± 0.040 33.35 ± 0.042 33.33 ± 0.042 33.17 ± 0.042		
	20	28.94 ± 0.031		29.54 ± 0.029 29.78 \pm 0.040 29.88 \pm 0.038 29.86 \pm 0.028 29.51 \pm 0.027						29.77 ± 0.027 29.79 \pm 0.027 29.70 \pm 0.026 29.73 \pm 0.024 29.34 \pm 0.033		
	60	$40\quad 25.31 \pm 0.050$ 22.99 ± 0.048		25.50 ± 0.057 25.67 ± 0.041 26.12 ± 0.049 26.11 ± 0.054 25.97 ± 0.055 22.95 ± 0.091 23.18 ± 0.061 23.80 ± 0.067 23.80 ± 0.071 23.78 ± 0.075						26.08 ± 0.054 26.08 ± 0.054 26.04 ± 0.056 26.05 ± 0.054 25.95 ± 0.055 23.81 ± 0.070 23.81 ± 0.070 23.79 ± 0.074 23.80 ± 0.073 23.78 ± 0.075		
barbara	10	32.93 ± 0.026		33.50 ± 0.018 33.66 ± 0.020 33.70 ± 0.021 33.66 ± 0.022 33.53 ± 0.020						33.72 ± 0.017 33.74 ± 0.017 33.69 ± 0.017 33.66 ± 0.017 33.44 ± 0.017		
	20	29.36 ± 0.032		29.83 ± 0.029 29.96 \pm 0.032 30.23 \pm 0.030 30.27 \pm 0.028 30.04 \pm 0.029						30.27 ± 0.029 30.27 ± 0.028 30.19 ± 0.026 30.24 ± 0.027 30.00 ± 0.030		
	40	25.68 ± 0.047		25.78 ± 0.048 25.79 ± 0.047 26.46 ± 0.043 26.51 ± 0.040 26.51 ± 0.039						26.52 ± 0.040 26.52 ± 0.040 26.51 ± 0.039 26.51 ± 0.040 26.51 ± 0.039		
boat	60 10	23.50 ± 0.032 31.78 ± 0.015		23.17 ± 0.039 23.57 ± 0.034 24.13 ± 0.037 24.15 ± 0.035 24.16 ± 0.035 32.73 ± 0.019 32.81 ± 0.018 32.82 ± 0.017 32.72 ± 0.015 32.61 ± 0.016						24.15 ± 0.036 24.15 ± 0.036 24.16 ± 0.035 24.16 ± 0.035 24.16 ± 0.035 32.73 ± 0.018 32.75 ± 0.018 32.72 ± 0.017 32.65 ± 0.017 32.49 ± 0.017		
	20	28.40 ± 0.017		29.14 ± 0.017 29.23 ± 0.019 29.37 ± 0.015 29.34 ± 0.015 29.05 ± 0.015						29.30 ± 0.018 29.30 ± 0.018 29.25 ± 0.017 29.27 ± 0.017 28.95 ± 0.018		
		40 21.95 \pm 0.053		25.45 ± 0.021 25.45 ± 0.021 26.01 ± 0.016 25.99 ± 0.016 25.92 ± 0.014						25.98 ± 0.012 25.98 ± 0.012 25.95 ± 0.013 25.96 ± 0.012 25.92 ± 0.014		
	60	23.64 ± 0.025		23.11 ± 0.028 23.72 ± 0.026 24.01 ± 0.026 24.01 ± 0.025 23.99 ± 0.025						24.01 ± 0.025 24.01 ± 0.025 24.00 ± 0.025 24.00 ± 0.025 23.99 ± 0.025		
hill	10	31.87 ± 0.029		32.63 ± 0.020 32.67 ± 0.019 32.71 ± 0.018 32.61 ± 0.018 32.47 ± 0.014						32.67 ± 0.016 32.67 ± 0.016 32.64 ± 0.016 32.55 ± 0.015 32.34 ± 0.013		
	20 40	28.82 ± 0.022 25.91 ± 0.022		29.23 ± 0.031 29.23 ± 0.031 29.48 \pm 0.024 29.48 \pm 0.023 29.29 \pm 0.023 25.70 ± 0.026 25.98 ± 0.024 26.36 ± 0.024 26.38 ± 0.022 26.37 ± 0.022						29.45 ± 0.021 29.45 ± 0.021 29.41 ± 0.023 29.42 ± 0.022 29.25 ± 0.024 26.38 ± 0.022 26.38 ± 0.022 26.37 ± 0.022 26.38 ± 0.022 26.37 ± 0.022		
	60	24.25 ± 0.017		23.45 ± 0.013 24.32 ± 0.018 24.60 ± 0.024 24.59 ± 0.024 24.60 ± 0.024						24.59 ± 0.022 24.59 ± 0.022 24.60 ± 0.024 24.60 ± 0.024 24.60 ± 0.024		
couple	10	31.80 ± 0.014		32.76 ± 0.009 32.81 ± 0.009 32.85 ± 0.010 32.80 ± 0.011 32.71 ± 0.011						32.76 ± 0.013 32.77 ± 0.013 32.75 ± 0.012 32.72 ± 0.011 32.59 ± 0.012		
	20	28.14 ± 0.023		28.93 ± 0.023 28.93 ± 0.023 29.16 ± 0.028 29.16 ± 0.029 28.86 ± 0.030						29.11 ± 0.028 29.11 \pm 0.028 29.07 \pm 0.029 29.08 \pm 0.029 28.76 \pm 0.030		
	60	$40\quad 24.93 \pm 0.035$ 23.25 ± 0.037		25.03 ± 0.026 25.05 ± 0.026 25.49 ± 0.026 25.50 ± 0.030 25.44 ± 0.032 22.76 ± 0.043 23.29 ± 0.036 23.59 ± 0.045 23.60 ± 0.044 23.59 ± 0.044						25.48 ± 0.031 25.48 ± 0.031 25.47 ± 0.030 25.47 ± 0.030 25.43 ± 0.032 23.60 ± 0.044 23.60 ± 0.044 23.59 ± 0.044 23.60 ± 0.044 23.59 ± 0.044		
fingerprint	10	30.27 ± 0.017		30.87 ± 0.015 30.87 ± 0.016 30.84 ± 0.016 30.80 ± 0.017 30.57 ± 0.016						30.88 ± 0.018 30.88 ± 0.019 30.83 ± 0.019 30.81 ± 0.020 30.50 ± 0.018		
		$20\ \ 26.64 \pm 0.010$		27.06 ± 0.014 27.06 ± 0.014 27.04 ± 0.014 27.12 ± 0.012 26.72 ± 0.013						27.10 ± 0.012 27.10 ± 0.012 26.93 ± 0.013 27.05 ± 0.013 26.70 ± 0.012		
	40	23.20 ± 0.018		23.68 ± 0.024 23.68 ± 0.024 23.96 ± 0.023 24.06 ± 0.022 24.05 ± 0.023						24.05 ± 0.022 24.05 ± 0.022 24.05 ± 0.022 24.05 ± 0.022 24.05 ± 0.023		
MRI	60 10	20.93 ± 0.034 40.06 ± 0.043		21.44 ± 0.029 21.44 \pm 0.029 21.85 \pm 0.041 21.97 \pm 0.037 21.98 \pm 0.037 39.19 ± 0.040 40.81 ± 0.033 40.89 ± 0.032 40.83 ± 0.034 40.71 ± 0.029						21.98 ± 0.037 21.98 ± 0.037 21.98 ± 0.037 21.98 ± 0.037 21.98 ± 0.037 40.79 ± 0.040 40.85 ± 0.038 40.83 ± 0.037 40.81 ± 0.038 40.61 ± 0.032		
	20	36.14 ± 0.067		34.47 ± 0.047 36.57 ± 0.062 36.70 ± 0.063 36.74 ± 0.065 36.60 ± 0.068						36.74 ± 0.063 36.77 ± 0.064 36.64 ± 0.068 36.72 ± 0.066 36.59 ± 0.067		
	40	32.22 ± 0.067		29.33 ± 0.055 32.31 \pm 0.071 32.53 \pm 0.069 32.53 \pm 0.069 32.52 \pm 0.069						32.49 ± 0.064 32.54 ± 0.070 32.52 ± 0.069 32.54 ± 0.069 32.52 ± 0.069		
	60	29.57 ± 0.056		26.13 ± 0.058 29.68 \pm 0.057 29.88 \pm 0.056 29.88 \pm 0.056 29.88 \pm 0.056						29.76 ± 0.060 29.86 ± 0.059 29.88 ± 0.056 29.88 ± 0.056 29.88 ± 0.056		
abbey	10	29.31 ± 0.035 $20\quad 25.53 \pm 0.034$		29.96 ± 0.029 29.96 ± 0.029 29.86 ± 0.031 29.87 ± 0.033 29.38 ± 0.030 25.91 ± 0.036 25.91 ± 0.036 25.87 ± 0.033 25.80 ± 0.034 25.27 ± 0.034						29.92 ± 0.028 29.92 ± 0.028 29.83 ± 0.032 29.83 ± 0.033 29.34 ± 0.030 25.87 ± 0.034 25.87 ± 0.034 25.72 ± 0.032 25.71 ± 0.034 25.16 ± 0.033		
	40	22.85 ± 0.036		22.94 ± 0.031 22.94 ± 0.031 23.10 ± 0.030 23.14 ± 0.024 23.10 ± 0.024						23.13 ± 0.025 23.13 ± 0.025 23.12 ± 0.025 23.12 ± 0.025 23.10 ± 0.025		
	60	21.60 ± 0.034		21.30 ± 0.031 21.58 ± 0.032 21.83 ± 0.038 21.85 ± 0.039 21.85 ± 0.039						21.85 ± 0.039		
airplane	10	31.42 ± 0.077		32.47 ± 0.088 32.54 \pm 0.086 32.49 \pm 0.085 32.50 \pm 0.084 32.05 \pm 0.082						32.36 ± 0.078 32.39 ± 0.077 32.30 ± 0.085 32.33 ± 0.084 31.78 ± 0.082		
cabin	20 40	27.52 ± 0.107 24.62 ± 0.112		28.53 ± 0.110 28.60 ± 0.117 28.70 ± 0.130 28.58 ± 0.129 28.21 ± 0.119 24.86 ± 0.124 24.91 ± 0.105 25.26 ± 0.113 25.28 ± 0.112 25.24 ± 0.121						28.57 ± 0.133 28.57 ± 0.133 28.45 ± 0.127 28.46 ± 0.127 28.11 ± 0.121 25.25 ± 0.120 25.25 ± 0.120 25.25 ± 0.118 25.25 ± 0.117 25.24 ± 0.120		
	60	22.89 ± 0.185		22.60 ± 0.176 22.97 ± 0.194 23.36 ± 0.206 23.39 ± 0.192 23.38 ± 0.193						23.38 ± 0.193		
airport	10	32.79 ± 0.042		33.41 ± 0.038 33.58 ± 0.036 33.51 ± 0.038 33.58 ± 0.043 33.09 ± 0.045						33.26 ± 0.041 33.34 ± 0.041 33.23 ± 0.044 33.37 ± 0.044 32.86 ± 0.040		
terminal		$20\ \ 28.74 \pm 0.071$		29.58 ± 0.050 29.72 \pm 0.044 29.95 \pm 0.055 29.97 \pm 0.054 29.67 \pm 0.049						29.70 ± 0.054 29.70 \pm 0.054 29.69 \pm 0.058 29.74 \pm 0.057 29.42 \pm 0.047		
	40 60	24.78 ± 0.060 22.72 ± 0.061		25.21 ± 0.055 25.21 ± 0.055 25.70 ± 0.077 25.71 ± 0.077 25.55 ± 0.070 22.59 ± 0.049 22.80 ± 0.063 23.22 ± 0.072 23.23 ± 0.073 23.20 ± 0.070						25.61 ± 0.074 25.61 ± 0.074 25.59 ± 0.074 25.60 ± 0.074 25.51 ± 0.076 23.21 ± 0.070 23.21 ± 0.070 23.21 ± 0.071 23.21 ± 0.071 23.20 ± 0.070		
allev	10	37.58 ± 0.075		37.65 ± 0.053 37.90 ± 0.070 38.44 ± 0.077 38.48 ± 0.076 38.42 ± 0.082						38.44 ± 0.079 38.44 ± 0.079 38.44 ± 0.079 38.45 ± 0.078 38.41 ± 0.081		
	20	33.90 ± 0.087		33.35 ± 0.087 34.15 ± 0.102 34.59 ± 0.101 34.59 ± 0.101 34.59 ± 0.101						34.59 ± 0.101 34.59 ± 0.101 34.60 ± 0.101 34.60 ± 0.101 34.59 ± 0.102		
		40 31.06 \pm 0.056		28.67 ± 0.123 31.12 ± 0.057 31.27 ± 0.058 31.26 ± 0.059 31.26 ± 0.060						31.16 ± 0.077 31.22 ± 0.064 31.26 ± 0.059 31.27 ± 0.060 31.26 ± 0.060		
amphitheater	60	29.65 ± 0.101		25.84 ± 0.110 29.68 ± 0.101 29.77 ± 0.102 29.76 ± 0.102 29.77 ± 0.102						29.46 ± 0.116 29.72 ± 0.102 29.77 ± 0.102 29.77 ± 0.103 29.77 ± 0.102		
	- 10	32.27 ± 0.038 $20\ \ 28.70 \pm 0.089$		32.87 ± 0.038 33.02 ± 0.034 32.94 ± 0.035 32.90 ± 0.036 32.39 ± 0.050 28.94 ± 0.061 29.09 ± 0.064 29.12 ± 0.071 29.02 ± 0.079 28.58 ± 0.078						32.94 ± 0.040 32.96 ± 0.037 32.83 ± 0.052 32.81 ± 0.049 32.23 ± 0.069 29.11 ± 0.073 29.11 ± 0.073 28.94 ± 0.081 28.91 ± 0.079 28.52 ± 0.077		
		40 25.88 \pm 0.135		25.40 ± 0.093 25.90 ± 0.126 26.02 ± 0.157 26.02 ± 0.173 25.99 ± 0.175						26.06 ± 0.158 26.06 ± 0.158 26.01 ± 0.173 26.01 ± 0.172 25.99 ± 0.175		
		$60, 24, 62 + 0, 077$		$23.45 + 0.159$ $24.64 + 0.080$ $24.72 + 0.081$ $24.72 + 0.082$ $24.72 + 0.082$						$24.72 + 0.096$ $24.73 + 0.093$ $24.72 + 0.082$ $24.72 + 0.082$ $24.72 + 0.082$		

C. Performance Studies With Bias-Variance Trade-Off

The bias-variance trade-off was investigated using many natural images. As shown in Fig. 1, a neighborhood size *B* was used to estimate p_i using the LJS method [33], and this was a significant factor when determining the bias. This tendency was also observed for the other different natural images, as illustrated in Fig. 4. Increasing *B* in the LJS method moved the bias-variance trade-off curves in the bottom right direction, meaning that the bias increased and the variance decreased. However, the role of the smoothing parameter *h* changed in the LJS method. Unlike in classical NLM method (see the NLM bias-variance curve in Fig. 1), increasing the smoothing parameter *h* beyond a certain point in the LJS method did not further decrease the variance in any of the natural images that

we tested. This is because increasing *h* will also increases the p_i values so that the resulting LJS estimator becomes closer to the noisy input image y_i due to the lack of bounds for the self-weights.

Our proposed methods (LMM-DB, LMM-RP) yielded trade-off curves that have decreased variances for increasing values of the smoothing parameter *h*. Figure 5 shows the trade-off curves for the cameraman example for different methods (LMM-DB, LMM-RP), different neighborhood sizes $(B = 2, 7)$, and different noise levels $(\sigma = 10, 40)$. Our proposed methods yielded bias-variance curves that were less than or equal to those in the LJS method for fixed *B* and σ . This tendency was also observed with other natural images, as illustrated in Fig. 4. It was important to choose appropriate

Fig. 4. Bias-variance curves for natural images using LJS+ [33] and our proposed LMM – DB^{one} and LMM – RP^{one} methods with a noise level of $\sigma = 10$. (a) *couple*. (b) *montage*. (c) *lena*. (d) *pepper*. (e) *house*. (f) *MRI*.

neighborhood sizes *B* in order for the LJS method to obtain a certain level of bias, but our proposed methods were able to achieve that same level of bias by adjusting the smoothing parameter *h*, which was the same as in classical NLM. Based on our results, it appears that the use of LMM-RP has slightly more advantages than using LMM-DB in terms of the PSNR, as shown in Table I, and the bias-variance trade-off curves, as shown in Fig. 5, for high noise levels.

D. Performance Studies With Visual Quality Assessment

The most important improvements in our proposed LMM-DB and LMM-RP methods when compared to the LJS method were achieved in terms of the visual quality. Figure 6 (a) shows the true cameraman image (left) and the noisy image (right) with a noise level of $\sigma = 10$. Figure 6 (b) presents the filtered images using the LJS method [33] with $B = 2$ and $B = 7$. Severe artifacts were observed in the background areas when using $B = 2$, and these artifacts were reduced when using $B = 7$. However, there were still some artifacts near the edges of objects. Our proposed LMM-DB and LMM-RP methods exhibited fewer image artifacts than were observed in the images processed using the LJS method for both $B = 2, 7$. This tendency was observed in many of the natural images, as shown in Fig. 7, especially in the high intensity flat areas. PSNR improvements in the LJS method were achieved with severe (when $B = 2$) or mild (when $B = 7$) artifacts; however, our proposed methods achieved both a high

Fig. 5. Bias-variance curves for LMM − DB^{one} and LMM – RP^{one} for comparison with LJS₊ for two neighborhood sizes $B = 2, 7$ and two noise levels $\sigma = 10, 40$. (a) LMM – DB^{one}($\sigma = 10$). (b) LMM – DB^{one}($\sigma = 40$). (c) LMM – $\text{RP}^{\text{one}}(\sigma = 10)$. (d) LMM – $\text{RP}^{\text{one}}(\sigma = 40)$.

PSNR and significantly reduced visual artifacts. This ability to reduce the number of visual artifacts in a denoised image is important in some applications, such as diagnostic medical imaging.

 (d)

Fig. 6. True, noisy ($\sigma = 10$), and filtered images using LJS₊ [33], and the proposed LMM – DB^{one} and LMM – RP^{one}. (a) True and noisy images ($\sigma = 10$). (b) LJS₊ [33]. (c) Proposed LMM – DB^{one}. (d) Proposed LMM − RPone.

E. Maximum Self-Weights: One vs. Stein's

Two maximum self-weights were proposed for use: the value one in (14) that was proposed in [4], and Stein's in (15) that was proposed in [32]. Figure 8 shows that the LMM − DBone method yielded an improved bias-variance curve and PSNR than did the $LMM - DB^{stein}$ method when the noise levels were low. For high noise levels $\sigma = 40$, the $LMM - DB^{stein}$ method yielded an improved PSNR and bias-variance curve than did the $LMM - DB^{one}$ method. However, these differences were not significant, as also illustrated in terms of the PSNR in Table I. In terms of the visual quality, no significant differences were observed between the two methods.

TABLE II PERCENTAGE (%) OF $c(\|\mathbf{s}\|)$ THAT EXCEED 2 USING LMM − DB AND LMM – RP METHODS, $\sigma = 10$, $B = 2$

		$\sigma = 10$	$\sigma = 20$	$\sigma = 40$	$\sigma = 60$
	LMM-DB one	0.32	0.04	0.03	0.05
cameraman	LMM DB ^{stein}	0.85	0.66	0.21	0.18
	LMM-DB one	0.00	0.00	0.00	0.00
fingerprint	LMM -DB ^{stein}	0.30	0.13	0.09	0.02
	LMM -DB one	0.10	0.05	0.10	0.13
MRI	LMM-DB ^{stein}	0.18	0.16	0.16	0.16
		$\sigma = 10$	$\sigma = 20$	$\sigma = 40$	$\sigma = 60$
	LMM-RP one	0.25	0.04	0.01	0.00
cameraman	LMM-RP ^{stein}	1.07	0.90	0.20	0.22
	LMM-RP one	0.01	0.00	0.00	0.00
fingerprint	LMM-RP ^{stein}	0.27	0.19	0.13	0.03
MRI	LMM -RP one	0.09	0.05	0.07	0.09

F. "Practical" Minimax Estimator

The proposed LMM-DB and LMM-RP methods are minimax with respect to $\|\mathbf{s}\| \leq Y_4$ and $\|\mathbf{s}\| \leq Y_3$, respectively, as shown in Fig. 3. However, these conditions impose upper bounds for the smoothing parameters *h* and the optimal *h*[∗], which means that the smoothing parameter values that yield the best PSNR may not be achievable. We empirically investigated this issue using many natural images.

Table II shows the ratio (percentage unit) of the number of pixels for which $c(\|\mathbf{s}\|) > 2$ to the total number of pixels in the cameraman, fingerprint, and MRI images when the optimal *h*[∗] for the highest PSNR was chosen based on the true images for the proposed LMM-DB and LMM-RP methods. For most of the pixels, the LMM-DB and LMM-RP values were minimax. The relationship between the percentage of pixels with $c(\|\mathbf{s}\|) > 2$ and the root mean squared error (RMSE) is illustrated in Fig. 9 for the cameraman and MRI images. Surprisingly, the optimal global smoothing parameters *h* for the lowest RMSE point (or the highest PSNR) of the LMM-DB and LMM-RP methods are very close to the smoothing parameters *h* such that the percentage of $c(\|\mathbf{s}\|) > 2$ is 0.1%. This phenomenon was not only observed in these two images. As shown in Table III, the pixel percentage of $c(\Vert s \Vert) > 2$ that do not require knowledge of the true image can still determine smoothing parameters that are able to yield comparable PSNR values to the best PSNR values obtained by using the optimal smoothing parameters calculated based on knowledge of the true image. This was observed in all of the natural images used in our simulations, with different noise levels, and when $B = 2$ was used. However, the criteria of using the pixel percentage of $c(||s||) > 2$ did not work very well for $B = 7$ in our simulations. These criteria can be potentially used when choosing a global smoothing parameter with our proposed methods as a heuristic approach without knowing the true image.

G. Computation Time for Algorithms

Table IV reports the computation time of the proposed methods in comparison with the classical NLM and LJS_{+} .

Fig. 7. Filtered results using LJS₊ [33] and the proposed LJS – RP^{one} method with a noise level of $\sigma = 10$ and neighborhood size B = 2. (a) *couple*. (b) *montage*. (c) *lena*. (d) *pepper*. (e) *house*. (f) *MRI*.

Fig. 8. Bias-variance curves and PSNR vs. varying neighborhood sizes (*B*) using classical NLM (only in the PSNR figure), LJS, and the proposed $LMM - DB^{stein}$ vs. $LMM - DB^{one}$ for the cameraman example.

We used 8 threads (Intel Core i7 2.8 GHz) when computing the patch distances for all methods. The local block size was $B = 2$, the patch size was 7×7 , and the window size was 31×31 . All parameters were fixed for all of results presented in this section. Adjusting these parameters can greatly reduce the running time. For example, setting $B = 4$, the patch size to 5×5 , and the window size to 13×13 reduces the computation time of the proposed methods to 0.60, 1.12, and 2.91 seconds (s) for 128^2 , 256^2 , and 512^2 images, respectively. However, analytically, the classical NLM requires $3|P||\Omega|$ + $4|\Omega| - 1$ operations per pixel where $|\Omega|$ is the number of elements in Ω_i and LJS₊ requires $3|\mathbf{P}||\Omega| + 4|\Omega| + 3|\mathbf{B}| + 5$ operations per pixel. It is reported in [33] that the additional operations for LJS_+ (3|**B**| + 6 operations) were negligible compared to the NLM filtering computation $(3|\mathbf{P}||\Omega|+4|\Omega|-$ 1 operations). Analytically, the additional computation for LMM – DB and LMM – RP is $3|\mathbf{B}| + 7$ operations, which is almost the same as the additional computation for $LJS₊$. Therefore, further implementation optimization is possible by exploiting the redundant computation of the patch distances for the minimax estimator and NLM weights.

VI. DISCUSSION

The classical NLM method was a significant work in image denoising [4], and required the determination of two important parameters for good denoising performance: a smoothing parameter and a self-weight value. The LJS method proposed

TABLE III

THE PSNR VALUES (dB) OF THE PROPOSED METHODS WITH $B = 2$ WHEN CHOOSING THE SMOOTHING PARAMETER SO AS TO YIELD THE HIGHEST PSNR USING THE TRUE IMAGE (TRUE), AND WHEN CHOOSING THE SMOOTHING PARAMETER SO AS TO YIELD THE PERCENTAGE OF $c(\|\mathbf{s}\|) > 2$ TO BE 0.1% (ESTIMATED) FOR DIFFERENT NOISE LEVELS

by Wu *et al.* [33] developed a state-of-the-art method for selfweight determination using JS estimation [37] and yielded superior results in terms of the PSNR compared to the other existing methods. However, since the LJS method did not impose an upper bound for self-weight estimation, the bias could no longer be controlled by the smoothing parameter, which resulted in visual quality degradation. Our proposed methods based on the Baranchik's minimax theorem [34] yielded comparable PSNR results to the state-of-the-art LJS method. By imposing upper bounds for the self-weights,

Fig. 9. Comparison plots of the RMSE vs. the smoothing parameter *h* and the percentage of $c(||s||) > 2$ vs. the same smoothing parameter when using LMM-DB and LMM-RP with B = 2 and $\sigma = 10$. (a) cameraman LMM – DB^{one}. (b) MRI LMM – DB^{one}. (c) cameraman LMM – RP^{one}. (d) MRI LMM $-$ RP^{one}.

TABLE IV EXECUTION TIME (S) COMPARISON. THIS WILL VARY WITH PARAMETER SELECTION

Image size	Classical NLM	$LJS+$		LMM-DB one LMM-RP one
$128*128$	0.65	0.89	0.90	0.90
$256*256$	1.57	2.37	2.39	2.38
512*512	4.90	7 14	719	7.18

the bias-variance trade-off was able to be controlled by a smoothing parameter, and substantial visual artifact reduction was achieved.

The focus of this article was self-weight parameter selection in the classical NLM filter with theoretical justification. As discussed in the Introduction, there are other factors that affect the performance of NLM based filters, and we expect that our proposed methods would not be able to achieve state-of-the-art denoising performance if there were no other optimizations performed except the self-weights. Indeed, our proposed methods with one patch size (non-adaptive neighborhood) and one global smoothing parameter were not able to achieve the level of denoising performance of the state-of-theart denoising methods such as BM3D [6]. However, when our proposed methods have incorporated some of the other factors into the NLM filters, such as local smoothing parameters and adaptive neighborhoods [10], they have great potential to achieve significantly improved denoising performance.

The minimax property of our proposed methods depends on the choice of smoothing parameters. When using sufficiently small smoothing parameters, the LMM-DB and LMM-RP methods are "practically" minimax according to Baranchik's theorem [34]. However, when large smoothing parameters are used, there may be some pixels that are not minimax for selfweight estimation. More empirical investigation showed that the optimal global smoothing parameter *h* that yielded the best PSNR only resulted in a very small portion of the pixels that did not have minimax self-weight estimators. In fact, this can be used as a useful heuristic when choosing a good smoothing parameter since testing the minimax properties of our proposed methods does not require the true image. More theoretical analysis for this observation, or a statistical analysis using many natural images as shown in [41], are potential extensions of this work. Therefore, our proposed methods do not only provide an optimal way to determine self-weights, but also provide a heuristic way to determine a good smoothing parameter.

VII. CONCLUSION

We proposed two methods, LMM-DB, LMM-RP, to determine the self-weights of NLM filters that are "practically" minimax, and this methods yielded a comparable PSNR, better bias-variance trade-offs, and reduced visual quality artifacts when compared to the results obtained using the state-of-theart LJS method. Our methods also provide a potentially useful heuristic way to determine a global smoothing parameter without knowledge of the original image.

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Minh Phuong Nguyen received the B.S.E. degree in biomedical electronics from the School of Electronics and Telecommunications, Hanoi University of Science and Technology, Hanoi, Vietnam, in 2013, and the M.S. degree in electrical engineering from the Ulsan National Institute of Science and Technology, Ulsan, South Korea. He is currently pursuing the Ph.D. degree with the Delft University of Technology, The Netherlands. His research interests include image reconstruction and image filtering for SPECT, image registration between histology images

and MRI, and electrocardiogram signal processing for authentication tasks.

Se Young Chun (M'10) received the B.S.E. degree from the School of Electrical Engineering, Seoul National University, Seoul, South Korea, in 1999, and the M.S.E. degree in electrical engineering systems, the M.S. degree in mathematics, and the Ph.D. degree in electrical engineering: systems from the University of Michigan, Ann Arbor, MI, USA, in 2005 and 2009, respectively.

From 2003 to 2005, he was a recipient of the two year scholarship from the Ministry of Information and Communications, South Korea. From 2009 to

2011, he was a Research Fellow with Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA, and from 2011 to 2013, he was a Research Fellow with Electrical Engineering and Computer Science and Radiology, University of Michigan. Since 2013, he has been an Assistant Professor with the School of Electrical and Computer Engineering, Ulsan National Institute of Science and Technology, Ulsan, South Korea. His research interests include image reconstruction, image denoising, image registration for multi-modal imaging systems, motion compensated image reconstruction, and statistical bio signal processing and biometrics.

Dr. Chun was a recipient of the 2010 Society of Nuclear Medicine Computer and Instrumentation Young Investigator Award, 2nd place. He was also a recipient of the 2015 Bruce Hasegawa Young Investigator Medical Imaging Science Award from the IEEE Nuclear and Plasma Sciences Society. He is a member of IEEE Computational Imaging Special Interest Group and an Editorial Board Member of *Nuclear Medicine and Molecular Imaging* (Springer).