Bounded Self-Weights Estimation Method for Non-Local Means Image Denoising Using Minimax Estimators

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Abstract-A non-local means (NLM) filter is a weighted average of a large number of non-local pixels with various image intensity values. The NLM filters have been shown to have powerful denoising performance, excellent detail preservation by averaging many noisy pixels, and using appropriate values for the weights, respectively. The NLM weights between two different pixels are determined based on the similarities between two patches that surround these pixels and a smoothing parameter. Another important factor that influences the denoising performance is the self-weight values for the same pixel. The recently introduced local James-Stein type center pixel weight estimation method (LJS) outperforms other existing methods when determining the contribution of the center pixels in the NLM filter. However, the LJS method may result in excessively large self-weight estimates since no upper bound is assumed, and the method uses a relatively large local area for estimating the self-weights, which may lead to a strong bias. In this paper, we investigated these issues in the LJS method, and then propose a novel local self-weight estimation methods using direct bounds (LMM-DB) and reparametrization (LMM-RP) based on the Baranchik's minimax estimator. Both the LMM-DB and LMM-RP methods were evaluated using a wide range of natural images and a clinical MRI image together with the various levels of additive Gaussian noise. Our proposed parameter selection methods yielded an improved bias-variance trade-off, a higher peak signal-to-noise (PSNR) ratio, and fewer visual artifacts when compared with the results of the classical NLM and LJS methods. Our proposed methods also provide a heuristic way to select a suitable global smoothing parameters that can yield PSNR values that are close to the optimal values.

Index Terms—James-Stein estimator, minimax estimator, non-local means, center pixel weight, bounded self-weight, image denoising.

I. INTRODUCTION

TMAGE denoising is a fundamental task in image processing, low-level computer vision, and medical imaging algorithms. The goal of denoising is to suppress image noise

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when restoring desired details using prior information about the images. For example, based on prior information regarding "smooth images", a simple filter, such as a Gaussian filter, can be designed as a weighted average of the image intensities of the pixels in the local neighborhood with non-adaptive weights. However, this type of filter blurs the edges and details of images because these features are not captured in the assumed prior information. Many edge-preserving denoising methods have been proposed, including bilaterial filters [1], [2], anisotropic diffusion [3], non-local means (NLM) filters [4], [5], collaborative filters (BM3D) [6], and total variation filters [7]. Many filters, including bilaterial filters, anisotropic diffusion, and NLM filters (but, not BM3D, see [8]), can be represented as the weighted averages of adaptive weights or adaptive smoothing [9]. It should be noted that it is important to select appropriate weights in these types of filters in order to obtain improved denoised image quality [8].

Classical NLM filters use the similarities between two local patches in a noisy image to determine the weights in nonlocal adaptive smoothing [4]. The NLM weights are obtained by first calculating the Euclidean distance between the two local patches, which is denoted d, and then by evaluating $\exp(-d^2/h^2)$, where h is a smoothing parameter. This method allows higher weights to be assigned to pixels with similar patches so that edges and details can be preserved through non-local weighted averaging.

There are four different factors that determine the output image quality of a NLM filter in terms of weights. 1) The first factor is the similarity measure d. The Euclidean distance is a usual choice [4], but other similarity measures have also been proposed, such as hypothesis testing with adaptive neighborhoods [10], principal component analysis (or the subspace based method) [11], [12], blockwise aggregation [13], rotationinvariant measures [14]-[16], shape-adaptive patches [17], and patch-based similarities with adaptive neighborhoods [18]. In multimodal medical imaging, inaccurate weights for noisy molecular images were enhanced by using additional high quality anatomical images [19], [20]. 2) The second factor is the strategy for determining the smoothing parameter h. Optimization strategies have been developed based on Stein's unbiased risk estimation (SURE) method for NLM with Gaussian noise [21], [22], NLM with Poisson noise [23], and blockwise NLM with Gaussian noise [24]. 3) The third factor is in selecting the function to use to determine the weights, such as $exp(-x^2)$. Other functions have also been proposed to calculate the weights, such as compact support func-

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tions [25], [26] and statistical distance functions [27], [28]. 4) The last factor, which is the focus of this article, is in how best to determine the self-weights for the same pixel in the input and output images.

The NLM weights for two different pixels are essentially determined by the distance between the two noisy local patches around these pixels. However, the weights for the same pixel, or the self-weights, are not affected by the noise in the patches and the distance is always 0. For an extremely noisy image, the self-weights will be relatively too large when compared to the other weights, which will cause the filter output to be almost the same as the input noisy image. Therefore, the use of appropriate self-weight values can significantly affect the quality of the denoised image. Many researchers have investigated strategies for determining the self-weights, which are also known as center pixel weights, in order to alleviate the so-called "rare patch effect." For the classical NLM filter proposed by Buades et al., the selfweights were set to be either one or the maximum weight in a neighborhood [4]. This strategy guaranteed that at least one or two of the largest weights would be the same. Doré and Cheriet also used the maximum weight in a neighborhood as the self-weight, but only if that maximum weight was large [29]. Brox and Cremers proposed a method to have at least n number of the weights to be the same [30], and Zimmer et al. considered the self-weight to be a free parameter during the estimation process [31]. Salmon developed a SURE-based method for determining the self-weights that accounted for the noise [32].

Recently, Wu et al. proposed a method to determine the self-weights using a James-Stein (JS) type estimator [33]. The idea of that work was to use a JS estimator to determine the reparametrized self-weight in a local neighborhood (called the local JS estimator (LJS)). The LJS method yielded the best peak signal-to-noise ratio (PSNR) results when compared to other existing self-weight selection methods [4], [32]. However, the method had some limitations. First, the LJS could yield self-weights that were theoretically much larger than 1 because no upper bound for the self-weights was assumed, and this may lead to severe rare patch artifacts. The JS estimator does not guarantee its optimality for bounded shrinkage parameters. Second, the original LJS method was tested with a relatively large local neighborhood when determining a self-weight because it was assumed that the selfweights were the same in the local neighborhood. However, the problem is that the selection of a local neighborhood size that is too large may introduce a strong bias into the resulting denoised images.

In this article, we investigate the original LJS method in terms of the local neighborhood size for self-weight estimation and the potential for excessive self-weight estimation when no upper bound is applied on the self-weight. We then propose novel self-weight estimation methods for NLM that account for bounded self-weights using Baranchik's minimax estimator [34], called local minimax self-weight estimation with direct bound (LMM-DB) and with reparametrization (LMM-RP). We evaluated our proposed methods using performance criteria including PSNR, the bias-variance tradeoff curve and visual quality assessment with a wide range of natural images and a real patient MRI image with various noise levels. We compared the performance of our proposed methods with a classical NLM filter using self-weights of 1 [4] and the state-of-the-art LJS method, which has already been shown to be the best among all other previous self-weight determination methods [33].

This article is an extension of a work that was presented at the 2016 IEEE International Symposium on Biomedical Imaging (ISBI) [35], and goes into more depth regarding the theory of the minimax estimator and provides an evaluation of the methods using a significantly larger image dataset.

This paper is organized as follows. Section II reviews the classical NLM filter and revisits the LJS method. Section III investigates the LJS method in terms of the local neighborhood size for self-weight estimation and the potential for excessively large self-weight estimates. Then, Section IV proposes novel LMM-DB and LMM-RP methods using Baranchik's minimax estimator in order to overcome two limitations of the LJS method. Section V illustrates the performance of our proposed methods by providing our simulation results. Lastly, Sections VI and VII discuss and then conclude this paper, respectively.

II. REVIEW OF THE LOCAL JAMES-STEIN SELF-WEIGHT ESTIMATION METHOD FOR THE NLM FILTER

In this section, we will briefly review both the classical NLM method proposed by Buades *et al.* [4] and the LJS self-weight selection method proposed by Wu *et al.* [33].

A. Reviewing the Classical Non-Local Means Filter

Let us assume that an image \mathbf{x} is contaminated by noise \mathbf{n} , which produces a noisy image \mathbf{y} :

$$\mathbf{y} = \mathbf{x} + \mathbf{n} \tag{1}$$

where **n** is zero-mean white Gaussian noise with standard deviation σ . The NLM filtered value at the pixel *i* is the weighted average of all pixels in a search region Ω_i :

$$\hat{x}_i = \frac{\sum_{j \in \Omega_i} w_{i,j} y_j}{\sum_{i \in \Omega_i} w_{i,j}}$$
(2)

where y_i is the *i*th element of \mathbf{y} , $w_{i,j}$ is the weight between the *i*th and *j*th pixels, and Ω_i is the set of all pixels in an area around the *i*th pixel, which could be an entire image. The similarity weight of the classical NLM is defined as:

$$w_{i,j} = \exp\left(\frac{-\left\|\mathbf{P}_{i}\mathbf{y} - \mathbf{P}_{j}\mathbf{y}\right\|^{2}}{2\left|\mathbf{P}\right|h^{2}}\right)$$
(3)

where \mathbf{P}_i is an operator used to extract a square-shaped patch centered at the *i*th pixel, $\|\cdot\|$ is an l_2 norm, $|\mathbf{P}|$ is the number of pixels within a patch, and *h* is a global smoothing parameter. Equation (3) implies that the self-weights $w_{i,i}$ are always equal to 1. Previous works on self-weights have shown that good strategies for determining the self-weights also affect the image quality of the NLM filtering [4], [29], [32], [33].

B. Reviewing Local James-Stein Self-Weight Estimation

The LJS method was proposed in order to determine $w_{i,i}$ as follows [33]. First, (2) was decomposed into two terms:

$$\hat{x}_{i} = \frac{W_{i}}{W_{i} + w_{i,i}} \hat{z}_{i} + \frac{w_{i,i}}{W_{i} + w_{i,i}} y_{i}$$
(4)

where $W_i = \sum_{j \in \Omega_i \setminus \{i\}} w_{i,j}$ and

$$\hat{z}_i = \sum_{j \in \Omega_i \setminus \{i\}} w_{i,j} y_j / W_i.$$
(5)

The terms \hat{z}_i do not contain $w_{i,i}$. Then, the LJS method reparametrized (4) using

$$p_i = \frac{w_{i,i}}{W_i + w_{i,i}} \tag{6}$$

so that (4) became:

$$\hat{x}_i = (1 - p_i)\,\hat{z}_i + p_i\,y_i.$$
 (7)

The problem of estimating the self-weights $w_{i,i}$ became the problem of estimating p_i . Lastly, the JS estimator [36], [37] for p_i was proposed:

$$p_i^{\text{LJS}} = 1 - \frac{\left(|\mathbf{B}| - 2\right)\sigma^2}{\left\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\right\|^2}$$
(8)

where \mathbf{B}_i is an operator used to extract a square-shaped neighborhood centered at the *i*th pixel, $|\mathbf{B}|$ is the number of pixels within that neighborhood, and σ is the known noise level.

Equation (8) implies that $p_i^{\text{LJS}} \in (-\infty, 1]$. Since the weights are non-negative, it was proposed to use the zero-lower bound for p_i^{LJS} as follows [33]:

$$\hat{x}_{i}^{\text{LJS}_{+}} = \left(1 - p_{i}^{\text{LJS}_{+}}\right)\hat{z}_{i} + p_{i}^{\text{LJS}_{+}}y_{i}$$
(9)

where

$$p_{i}^{\text{LJS}_{+}} := [p_{i}^{\text{LJS}}]_{+} = \left[1 - \frac{(|\mathbf{B}| - 2)\sigma^{2}}{\|\mathbf{B}_{i}\mathbf{y} - \mathbf{B}_{i}\hat{\mathbf{z}}\|^{2}}\right]_{+}$$
(10)

and $[s]_+ := \max(s, 0)$. Wu *et al.* also mentioned that a userdefined upper bound for p_i can be used, but did not investigate further [33]. It should be noted that the JS estimator does not guarantee its optimality when bounding p_i^{LJS} in (8).

III. LIMITATIONS OF THE LOCAL JAMES-STEIN SELF-WEIGHT ESTIMATION FOR THE NLM FILTER

We now investigate two limitations of the original LJS method [33] in terms of the size of local neighborhoods for self-weight estimation, and the potential for excessive self-weight estimation.

A. Size of Local Neighborhood for Self-Weight Estimation

In the method described in [33], there are two implicit steps required in order to obtain the LJS self-weight estimator (10). The first step is to choose a local set of pixels around the *i*th pixel, referred to as set Ω_i^B , that correspond to the operator **B**_i, and assume that:

$$\hat{x}_{j} = (1 - p_{i})\hat{z}_{j} + p_{i}y_{j}, \quad j \in \Omega_{i}^{B}.$$
 (11)



Fig. 1. Bias-variance curves (cameraman example) for the classical NLM and LJS methods (LJS₊) for different sizes of local neighborhoods (B). The curves were plotted while varying the smoothing parameter h ($log_2 h \in [1.8, 3.2]$).

Based on the works of Stein [36] and James and Stein [37], if $|\mathbf{B}| \ge 3$, then for a neighborhood Ω_i^B extracted using \mathbf{B}_i ,

$$\hat{x}_j = \left(1 - p_i^{\text{LJS}_+}\right)\hat{z}_j + p_i^{\text{LJS}_+}y_j, \quad j \in \Omega_i^B.$$
(12)

is a dominant estimator for x_j "locally" in Ω_i^B . The LJS method used the zero lower bound when estimating p_i in order to obtain a realistic non-negative self-weight value. This was also a good choice in terms of the estimator performance since the positive part of the JS estimator is dominant over the original JS estimator, according to the works of Baranchik [34], [38] and Efron and Morris [39].

The second implicit step is to assign the resulting p_i^{LJS} to p_i in (7) for only the single pixel *i* so that:

$$\hat{x}_i^{\text{LJS}} = \left(1 - p_i^{\text{LJS}}\right)\hat{z}_i + p_i^{\text{LJS}}y_i.$$
(13)

Wu *et al.* evaluated the LJS method with $|\mathbf{B}| = 15 \times 15$ [33], which seems relatively large.

Based on this implicit two-step interpretation, we can surmise that using a smaller size of $|\mathbf{B}|$ may be more desirable for obtaining a less biased estimate of p_i since the assumption of having the same p_i in Ω_i^B is less likely to be true for larger sizes of Ω_i^B . Figure 1 confirms our conjecture. The bias-variance curves of the LJS method yielded better biasvariance trade-offs than those in the classical NLM method for both large local neighborhoods with a half window size B = 7 $(|\mathbf{B}| = 15 \times 15)$ and small local neighborhoods with B = 2 $(|\mathbf{B}| = 5 \times 5)$. However, using larger local neighborhood sizes for estimating p_i yielded stronger biases than those estimated using smaller sizes for the same level of variance.

B. Excessively Large Self-Weight Estimates

In the LJS method for determining the self-weights by estimating values for p_i [33], it is theoretically possible that the self-weights have excessively high values. For example, (6) suggests that if $p_i = 1$ and $W_i > 0$, then $w_{i,i} \gg 1$. Slight artifacts were observed in [33] in the background area that were potentially caused by excessive self-weight estimates when a relatively larger neighborhood size $|\mathbf{B}| = 15 \times 15$ was used. We observed a significantly higher degree of degradation in the visual image quality in the background area when the



Fig. 2. Denoised image of the cameraman example using the original LJS method [33] with no upper bound for the self-weights (top left), estimated p_i values (top right), calculated W_i (bottom left), and resulting self-weights ($w_{i,i}$) showing excessive self-weights (bottom right). B = 2 and σ = 10.

size of $|\mathbf{B}|$ in (8) was small, as shown in the image in the top left figure of Fig. 2.

We investigated this issue using an example of the cameraman image that was denoised using the LJS method [33], but with a smaller neighborhood size $|\mathbf{B}| = 5 \times 5$. For areas with more details, such as edges and textures, large p_i values were estimated and yielded large self-weights, as shown in the top right figure of Fig. 2. However, since the values for W_i were also very small in these areas, as shown in the bottom left image in Fig. 2, the resulting self-weight map yielded values close to 1 in the areas with details as shown in the bottom right image in Fig. 2.

In contrast, for areas with almost no details, such as those with a flat intensity background, relatively smaller p_i values were estimated, some of which were much larger than 0 while the rest were closer to 0, as shown in the top right image in Fig. 2. However, since the W_i values for the flat areas were relatively large, as shown in the bottom left image in Fig. 2, some of the estimated p_i values obtained using the LJS method (LJS+) were estimated to yield excessively large self-weights that were much larger than 1, as shown in the bottom right image of Fig. 2. Consequently, these excessively large self-weights caused severe rare patch artifacts in the filtered image, which resulted in visual quality degradation, as observed in the top left image of Fig. 2.

IV. LOCAL MINIMAX ESTIMATION METHODS FOR UPPER BOUNDED SELF-WEIGHTS IN A NLM FILTER

In this section, we propose two local upper bounded selfweight estimation methods that use Baranchik's minimax estimator [34].

A. Bounded Self-Weights

It is usually assumed that the self-weights satisfy $w_{i,i} \in [0, 1]$. However, there are many possible upper bounds for the self-weights, including 1 [4] or some positive value that is

possibly less than 1 based on SURE [32]. In this article, two different upper bound values $w_{i,i}^{\max}$ for the self-weights were evaluated such that $0 \le w_{i,i} \le w_{i,i}^{\max}$. One upper bound was:

$$w_{i\,i}^{\max-\text{one}} = 1,\tag{14}$$

which is the usual choice for the self-weights in the classical NLM method [4]. The other upper bound was:

$$w_{i,i}^{\max-\text{stein}} = \exp\left(-\sigma^2/h^2\right),$$
 (15)

which was motivated by the SURE-based NLM self-weights [32]. We assume that σ is known and h is pre-determined, which means that the upper bound for the self-weights can also be determined in advance. Equation (15) takes the noise level into account. As σ is smaller, the maximum self-weight in (15) is closer to one. It should be noted that the difference between (15) and (14) will be greater at higher noise levels.

Since p_i is estimated instead of $w_{i,i}$, it is necessary to derive the range of p_i that corresponds to $0 \le w_{i,i} \le w_{i,i}^{\max}$. From (6), the derivative of p_i with respect to $w_{i,i}$ is nonnegative as follows:

$$\frac{d}{dw_{i,i}}p_i = \frac{W_i}{(W_i + w_{i,i})^2} \ge 0$$

since $W_i \ge 0$. Therefore, p_i is a non-decreasing function of $w_{i,i}$ and for $0 \le w_{i,i} \le w_{i,i}^{\max}$, the range of p_i will be

$$0 \le p_i \le \frac{w_{i,i}^{\max}}{W_i + w_{i,i}^{\max}} =: p_i^{\max} \le 1.$$

Note that if $W_i = 0$, then $p_i^{\max} = 1$. The estimator $p_i^{\text{LJS}_+}$ in (10) automatically guarantees that $0 \le p_i \le 1$ if $|\mathbf{B}| \ge 2$. However, since $W_i > 0$ generally holds for most real images with noise, it is necessary to constrain p_i to be less than or equal to the upper bound p_i^{\max} , which is usually less than one.

B. Local Minimax Self-Weight Estimation With Direct Bound

Enforcing the upper limit p_i^{max} on the estimated p_i in (10) using min $(p_i^{\text{LJS}+}, p_i^{\text{max}})$ breaks the optimality of the JS estimator if $p_i^{\text{max}} < 1$. In this article, we propose using Baranchik's minimax estimator [34] to incorporate bounded self-weights into the estimator (see Baranchik [34], Erfon and Morris [39], and Strawderman [40] for more details on this minimax estimator).

Theorem 1 (Baranchik): For $\mathbf{y} \sim \mathcal{N}_r(\mathbf{x}, \sigma^2 \mathbf{I}), r \geq 3$, and loss $L(\mathbf{x}, \hat{\mathbf{x}}) = \|\mathbf{x} - \hat{\mathbf{x}}\|$, an estimator of the form $\hat{\mathbf{x}} = q\mathbf{y}$ where

$$q = \left[1 - c\left(\|\mathbf{y}\|\right) \frac{\sigma^2(r-2)}{\|\mathbf{y}\|^2}\right]$$
(16)

is the minimax, provided that:

(i) $0 \le c (||\mathbf{y}||) \le 2$ and

(ii) the function $c(\cdot)$ is nondecreasing.

Here y shrinks toward 0 which is the initial estimate of x.



Fig. 3. Graphical illustrations of the original and positive part JS estimators without upper bounds, and the proposed minimax self-weight estimators with upper bounds in terms of $c (||\mathbf{s}||)$ vs. $||\mathbf{s}||$. (a) Original and positive-part JS estimators. (b) Proposed minimax estimators with bounds.

The original JS estimator and its positive part are special cases of Baranchik's minimax estimator. For the original JS estimator (8):

$$c\left(\|\mathbf{s}\|\right) = 1,\tag{17}$$

where $\mathbf{s} = \mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}$ so that both conditions (i, ii) of the Baranchik's theorem are satisfied. In the positive part estimator (9), it can be shown that:

$$c\left(\|\mathbf{s}\|\right) = \begin{cases} \frac{\|\mathbf{s}\|^2}{\sigma^2(r-2)}, & 0 \le \|\mathbf{s}\| \le Y_1\\ 1, & \text{otherwise} \end{cases}$$
(18)

where $Y_1 := \sigma \sqrt{r-2}$. The original and positive part JS estimators are illustrated in Fig. 3 (a).

We propose a new local minimax self-weight estimation method that uses a direct bound with a specific upper-bound value, as follows:

$$p_i^{\text{LMM}-\text{DB}} := \min(p_i^{\text{LJS}_+}, p_i^{\text{max}}).$$
(19)

This estimator is minimax under certain conditions that can be derived using Baranchik's minimax estimator theorem. According to this theorem, this operation can be interpreted as follows:

$$c\left(\|\mathbf{s}\|\right) = \begin{cases} \frac{\|\mathbf{s}\|^2}{\sigma^2(r-2)}, & 0 \le \|\mathbf{s}\| \le \mathbf{Y}_1\\ 1, & \mathbf{Y}_1 < \|\mathbf{s}\| \le \mathbf{Y}_2 & (20)\\ \frac{\|\mathbf{s}\|^2 \left(1 - p^{\max}\right)}{\sigma^2(r-2)}, & \mathbf{Y}_2 < \|\mathbf{s}\| \end{cases}$$

where $Y_2 := \sigma \sqrt{(r-2)/(1-p^{\max})}$. We call this a local minimax self-weight estimator using direct bound (LMM-DB), which is illustrated in Fig. 3 (b) where $Y_4 := \sigma \sqrt{2(r-2)/(1-p^{\max})}$. However, note that LMM-DB is not minimax for $||\mathbf{s}|| > \mathbf{Y}_4$. Fortunately, $||\mathbf{s}|| = ||\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}||$ can be limited by adjusting the smoothing parameter *h* by making it smaller so that all $||\mathbf{s}|| \le \mathbf{Y}_4$ and $c (||\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}||) \le 2$. Then, the LMM-DB becomes "practically" a minimax estimator. Let us denote the maximum *h* that satisfies $||\mathbf{s}|| \le \mathbf{Y}_4$ as h^{max} .

In this case, a question can be raised: will the optimal value for *h* fall into the range of *h* that satisfies $\|\mathbf{s}\| \leq Y_4$? Interestingly, our simulations with many natural images showed that the optimal smoothing parameter h^* based on the true images is very close to h^{max} . This is because the LMM-DB yielded $p^{\text{max}} \rightarrow 1$ so that $Y_2 \rightarrow \infty$, and almost all $\|\mathbf{B}_i\mathbf{y} - \mathbf{B}_i\hat{\mathbf{z}}\|$ were less than or equal to Y_4 . Therefore, $p_i^{\text{LMM-DB}}$ is "practically" a minimax value based on Baranchik's theorem for many natural images. Moreover, the LMM-DB method may provide a way to choose the optimal global smoothing parameter value *h* without knowing the underlying true image. We empirically investigate this issue in Section V.

C. Local Minimax Self-Weight Estimation With Re-Parametrization

The LMM-DB algorithm set p to be the same p^{max} for a wide range of $\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|$ values. We now propose another new method, called the local minimax self-weight estimation with reparametrization (LMM-RP) method, that assigns different p values for different $\|\mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}}\|$.

We reparametrized p_i in (7) in the following way:

$$\hat{x}_{i} = \hat{z}_{i}(p_{i}/p_{i}^{\max})p_{i}^{\max}(y_{i} - \hat{z}_{i}) = \hat{z}_{i} + p_{i}^{T}(y_{i}^{T} - \hat{z}_{i}^{T})$$
(21)

$$= (1 - p_i^{\max})\hat{z}_i + \hat{z}_i^{\mathrm{T}} + p_i^{\mathrm{T}}(y_i^{\mathrm{T}} - \hat{z}_i^{\mathrm{T}})$$
(22)

where $\hat{z}_i^{\mathrm{T}} = p_i^{\max} \hat{z}_i, \ y_i^{\mathrm{T}} = p_i^{\max} y_i$, and

$$p_i^{\rm T} = \frac{1}{p_i^{\rm max}} \frac{w_{i,i}}{W_i + w_{i,i}}.$$
 (23)

Note that for $0 \le w_{i,i} \le w_{i,i}^{\max}$, p_i^{T} is an increasing function of $w_{i,i}$ and the range of p_i^{T} is $0 \le p_i^{T} \le 1$. We propose to use the positive part of the JS estimator to estimate the reparametrized p_i^{T} , as follows:

$$p_{i}^{\mathrm{T,LJS}_{+}} = \left[1 - \frac{(|\mathbf{B}| - 2) (p_{i}^{\mathrm{max}})^{2} \sigma^{2}}{\|\mathbf{B}_{i} \mathbf{y}^{\mathrm{T}} - \mathbf{B}_{i} \hat{\mathbf{z}}^{\mathrm{T}}\|^{2}}\right]_{+}$$
$$= \left[1 - \frac{(|\mathbf{B}| - 2) \sigma^{2}}{\|\mathbf{B}_{i} \mathbf{y} - \mathbf{B}_{i} \hat{\mathbf{z}}\|^{2}}\right]_{+} = p_{i}^{\mathrm{LJS}_{+}}.$$
 (24)

This method is equivalent to using a multiplicative factor p_i^{max} for the original JS shrinkage (9):

$$\hat{x}_{i}^{\text{LMM}-\text{RP}} = (1 - p_{i}^{\text{LMM}-\text{RP}})\hat{z}_{i} + p_{i}^{\text{LMM}-\text{RP}}y_{i}$$
 (25)

where

$$p_i^{\text{LMM-RP}} = p_i^{\text{max}} \left[1 - \frac{\left(|\mathbf{B}| - 2 \right) \sigma^2}{\left\| \mathbf{B}_i \mathbf{y} - \mathbf{B}_i \hat{\mathbf{z}} \right\|^2} \right]_+.$$
 (26)

This proposed LMM-RP estimator is not dominant when estimating x_i , but rather is dominant when estimating $p_i^{\max}x_i$,

as shown in (22). Thus, the positive part JS estimator does not guarantee that the LMM-RP is dominant.

Baranchik's minimax estimation theorem can be used to analyze the LMM-RP estimator as follows:

$$c\left(\|\mathbf{s}\|\right) = \begin{cases} \frac{\|\mathbf{s}\|^2}{\sigma^2(r-2)}, & 0 \le \|\mathbf{s}\| \le Y_1 \\ \frac{\|\mathbf{s}\|^2(1-p^{\max})}{\sigma^2(r-2)} + p^{\max}, & Y_1 < \|\mathbf{s}\| \end{cases}$$
(27)

where if $\|\mathbf{s}\|$ is $Y_3 := \sigma \sqrt{(2 - p^{\max})(r - 2)/(1 - p^{\max})}$, then $c(\|\mathbf{s}\|) = 2$. The LMM-RP method is also illustrated in Fig. 3 (b), and is minimax if $\|\mathbf{s}\| \le Y_3$. The global smoothing parameter *h* can be adjusted so that this condition is satisfied for different images. As in the case of the LMM-DB, it turns out that the optimal global smoothing parameter h^* and the upper bound *h* that satisfies $\|\mathbf{s}\| \le Y_3$ are also very close to each other when the LMM-RP method is applied to many natural images. Therefore, the LMM-RP method is "practically" a minimax. The following table summarizes the LJS self-weight estimation methods.

Summary of Self-Weight Estimation Methods

$$\begin{aligned} \mathbf{LJS}_{+} & [33]: \\ p_{i}^{\text{LJS}_{+}} = \left[1 - (|\mathbf{B}| - 2) \sigma^{2} / \|\mathbf{B}_{i}\mathbf{y} - \mathbf{B}_{i}\hat{\mathbf{z}}\|^{2} \right]_{+} \\ \hat{x}_{i}^{\text{LJS}_{+}} &= (1 - p_{i}^{\text{LJS}_{+}})\hat{z}_{i} + p_{i}^{\text{LJS}_{+}} y_{i} \\ \mathbf{LMM} - \mathbf{DB}: \\ p_{i}^{\text{LMM}-\text{DB}} &= \min(p_{i}^{\text{LJS}_{+}}, p_{i}^{\text{max}}) \\ \hat{x}_{i}^{\text{LMM}-\text{DB}} &= (1 - p_{i}^{\text{LMM}-\text{DB}})\hat{z}_{i} + p_{i}^{\text{LMM}-\text{DB}} y_{i} \\ \mathbf{LMM} - \mathbf{RP}: \\ p_{i}^{\text{LMM}-\text{RP}} &= p_{i}^{\text{LJS}_{+}} p_{i}^{\text{max}} \\ \hat{x}_{i}^{\text{LMM}-\text{RP}} &= (1 - p_{i}^{\text{LMM}-\text{RP}})\hat{z}_{i} + p_{i}^{\text{LMM}-\text{RP}} y_{i} \end{aligned}$$

V. SIMULATION RESULTS

A. Simulation Setup

Ten natural images¹ (cameraman, lena, montage, house, pepper, barbara, boat, hill, couple, fingerprint) and five images from the SUN database² (abbey, airplane cabin, airport terminal, alley, amphitheater) were used in our study as noise-free images (128×128 , 256×256 , or 512×512 pixels, 8 bits). A real patient MRI (512×512 pixels, 8 bits) that was acquired and processed under institutional review board (IRB) approved protocols was also used. White Gaussian noise was added to each input image with various standard deviations $\sigma \in \{10, 20, 40, 60\}$.

All algorithms were implemented using MATLAB R2015b (The Mathworks, Inc., Natick, MA, USA). The patch size and search window size of the NLM filter were chosen to be 7×7 and 31×31 , respectively, which were the same

as those used in [33]. Both the state-of-the-art LJS algorithm and the proposed algorithms were tested using $B = 1, \dots, 9$ where $|\mathbf{B}| = (2B + 1)^2 > 3$.

The global smoothing parameter h was chosen empirically to yield the best PSNR:

PSNR
$$(\hat{\mathbf{x}}) = 10 \log_{10} \frac{255^2}{\|\hat{\mathbf{x}} - \mathbf{x}\|^2 / N},$$
 (28)

where N is the size of the image. In addition to the PSNR, the mean bias vs. the mean variance trade-off curves were used as performance measures for the different smoothing parameter values h:

$$\overline{\text{bias}^2} = \frac{1}{N} \sum_{i=1}^{N} (\bar{x}_i - x_i)^2,$$
 (29)

$$\overline{\text{var}} = \frac{1}{N} \sum_{i=1}^{N} \frac{1}{k-1} \sum_{j=1}^{k} \left(\hat{x}_{ij} - \bar{x}_i \right)^2, \quad (30)$$

where k is the number of realizations (20 in our simulation), \hat{x}_{ij} is the *j*th estimation at the *i*th pixel, and \bar{x}_i is the mean of \hat{x}_{ij} , as given by:

$$\bar{x}_i = \frac{1}{k} \sum_{j=1}^k \hat{x}_{ij}.$$

A visual quality assessment was also performed.

B. Performance Studies Using the PSNR

In order to estimate values of p_i for a fixed neighborhood size B, the optimal NLM smoothing parameter h^* was determined such that the PSNR was maximized. In our proposed methods, the two maximum self-weights in (14) and (15) were used. The LMM-DB and LMM-RP methods given by (14) are denoted LMM - DBone and LMM - RPone, while the LMM-DB and LMM-RP methods given by (15) are denoted LMM - DB^{stein} and LMM - RP^{stein}. Table I summarizes the quantitative PSNR results for the 16 images with 4 different noise levels. When B = 7, our proposed LMM-DB and LMM-RP methods based on Baranchik's minimax estimator yielded much better PSNR results than did setting the self-weight to one in the classical NLM method [4], and comparable PSNR values to the LJS method based on the JS estimator [33]. When B = 2, our proposed LMM-DB and LMM-RP methods yielded better PSNR values than did the LJS.

For the five examples of lena, house, peppers, barbara, boat with $\sigma = 20$, PSNRs of our proposed methods (global smoothing parameter and fixed neighborhoods, but adaptive self-weight) were $0.72 \sim 0.97$ dB better than classical NLM. In [10], it is reported that for the same five examples with the same level of noise, the work of Kervrann et al. (fixed self-weight, but local smoothing parameters and adaptive neighborhoods) yielded $0.99 \sim 1.55$ dB better PSNR than classical NLM. Self-weights, local smoothing parameter, neighborhoods size are important factors in the NLM filter to determine output image quality.

¹Available online at: http://www.cs.tut.fi/~foi/GCF-BM3D/BM3D_images. zip as the date of 16 Nov. 2015. ²Available online at: http://vision.princeton.edu/projects/2010/SUN/ as the

²Available online at: http://vision.princeton.edu/projects/2010/SUN/ as the date of 16 Sep. 2016.

TABLE I PSNR (dB) Summary (Mean \pm Standard Deviation) for Various Natural Images

		B = 2				$\mathbf{B} = 7$						
	σ	Classical NLM	LJS ₊	LMM-DB ^{one}	LMM-RP ^{one}	LMM-DB stein	LMM-RP stein	LJS ₊	LMM-DB ^{one}	LMM-RP ^{one}	LMM-DB stein	LMM-RP stein
cameraman	10	32.42 ± 0.034	33.12 ± 0.031	$\textbf{33.32} \pm \textbf{0.030}$	33.29 ± 0.029	33.17 ± 0.029	33.04 ± 0.030	32.98 ± 0.035	33.10 ± 0.036	33.05 ± 0.037	32.98 ± 0.038	32.85 ± 0.040
	20	28.48 ± 0.052	29.12 ± 0.052	29.46 ± 0.056	29.29 ± 0.056	29.27 ± 0.060	28.97 ± 0.052	29.32 ± 0.062	29.34 ± 0.059	29.04 ± 0.054	29.11 ± 0.058	28.80 ± 0.049
	40 60	25.35 ± 0.059 23.19 ± 0.065	25.39 ± 0.083 22.88 ± 0.062	23.89 ± 0.075 23.39 ± 0.055	20.11 ± 0.082 23 74 ± 0.065	26.08 ± 0.087 23.67 ± 0.069	23.74 ± 0.079 23.54 ± 0.061	25.98 ± 0.073 23.68 ± 0.068	25.98 ± 0.073 23.69 ± 0.068	25.94 ± 0.080 23.62 ± 0.071	23.96 ± 0.083 23.63 ± 0.070	23.63 ± 0.076 23.54 ± 0.059
lena	10	$\frac{23.19 \pm 0.005}{33.90 \pm 0.018}$	$\frac{22.88 \pm 0.002}{34.52 \pm 0.017}$	$\frac{23.39 \pm 0.033}{34.74 \pm 0.017}$	34.81 ± 0.003	34.77 ± 0.007	34.69 ± 0.001	$\frac{23.08 \pm 0.008}{34.82 \pm 0.020}$	34.83 ± 0.020	34.80 ± 0.019	34.76 ± 0.018	34.63 ± 0.018
	20	30.78 ± 0.031	30.90 ± 0.023	31.25 ± 0.033	31.50 ± 0.032	31.51 ± 0.032	31.31 ± 0.029	31.51 ± 0.031	31.51 ± 0.031	31.45 ± 0.028	31.48 ± 0.031	31.27 ± 0.029
	40	27.64 ± 0.032	26.94 ± 0.028	27.74 ± 0.033	28.08 ± 0.029	28.08 ± 0.029	28.06 ± 0.028	$\textbf{28.10} \pm \textbf{0.028}$	$\textbf{28.10} \pm \textbf{0.028}$	28.07 ± 0.028	28.08 ± 0.029	28.06 ± 0.028
montage	60	$\frac{25.60 \pm 0.052}{24.68 \pm 0.045}$	24.38 ± 0.040	25.66 ± 0.051	$\frac{26.02 \pm 0.052}{25.67 \pm 0.030}$	$\frac{26.01 \pm 0.053}{25.65 \pm 0.042}$	$\frac{26.02 \pm 0.054}{25.55 \pm 0.045}$	$\frac{26.01 \pm 0.051}{25.12 \pm 0.040}$	$\frac{26.01 \pm 0.051}{25.20 \pm 0.046}$	$\frac{26.02 \pm 0.054}{25.28 \pm 0.042}$	$\frac{26.02 \pm 0.053}{25.46 \pm 0.045}$	$\frac{26.02 \pm 0.054}{25.24 \pm 0.047}$
moniuge	20	34.08 ± 0.045 30 35 ± 0.088	35.19 ± 0.043 30.74 ± 0.062	35.60 ± 0.042 31.29 ± 0.067	35.07 ± 0.039 31.38 ± 0.070	35.05 ± 0.042 31 40 + 0 078	35.55 ± 0.045 31.06 ± 0.068	35.12 ± 0.049 31.00 ± 0.076	35.39 ± 0.046 31.13 ± 0.073	35.38 ± 0.042 31.07 ± 0.068	35.46 ± 0.045 31.18 ± 0.068	35.34 ± 0.047 30.81 ± 0.063
	40	26.24 ± 0.063	26.30 ± 0.072	26.98 ± 0.070	27.29 ± 0.063	27.30 ± 0.064	27.20 ± 0.061	27.01 ± 0.052	27.03 ± 0.052	27.04 ± 0.055	27.08 ± 0.053	27.00 ± 0.053
	60	23.76 ± 0.104	23.48 ± 0.116	24.16 ± 0.113	24.61 ± 0.115	24.60 ± 0.106	24.38 ± 0.115	24.33 ± 0.092	24.33 ± 0.092	24.36 ± 0.090	24.37 ± 0.088	24.16 ± 0.095
house	10	34.57 ± 0.038	35.02 ± 0.039	35.36 ± 0.041	35.38 ± 0.039	35.34 ± 0.043	35.29 ± 0.046	35.31 ± 0.042	35.32 ± 0.043	35.25 ± 0.044	35.21 ± 0.047	35.12 ± 0.045
	20	31.43 ± 0.063	31.54 ± 0.048	32.13 ± 0.050	32.39 ± 0.048	32.39 ± 0.050	32.19 ± 0.067	32.30 ± 0.052	32.31 ± 0.054	32.26 ± 0.058	32.30 ± 0.056	32.10 ± 0.068
	40	27.62 ± 0.044 25.01 ± 0.092	27.18 ± 0.038 24.24 ± 0.087	27.84 ± 0.049 25.17 ± 0.098	28.37 ± 0.037 25.65 ± 0.095	28.37 ± 0.039 25.65 ± 0.087	28.33 ± 0.043 25.65 ± 0.088	28.35 ± 0.041 25.65 ± 0.088	28.35 ± 0.041 25.65 ± 0.088	28.34 ± 0.043 25.65 ± 0.087	28.35 ± 0.042 25.66 ± 0.087	28.33 ± 0.043 25.65 ± 0.088
peppers	10	$\frac{25.01 \pm 0.052}{32.62 \pm 0.056}$	$\frac{21.21 \pm 0.007}{33.37 \pm 0.042}$	$\frac{23.17 \pm 0.098}{33.53 \pm 0.048}$	33.56 ± 0.049	33.49 ± 0.051	33.39 ± 0.050	$\frac{23.05 \pm 0.000}{33.28 \pm 0.043}$	33.37 ± 0.040	33.35 ± 0.042	33.33 ± 0.042	33.17 ± 0.042
	20	28.94 ± 0.031	29.54 ± 0.029	29.78 ± 0.040	$\textbf{29.88} \pm \textbf{0.038}$	29.86 ± 0.028	29.51 ± 0.027	29.77 ± 0.027	29.79 ± 0.027	29.70 ± 0.026	29.73 ± 0.024	29.34 ± 0.033
	40	25.31 ± 0.050	25.50 ± 0.057	25.67 ± 0.041	26.12 ± 0.049	26.11 ± 0.054	25.97 ± 0.055	26.08 ± 0.054	26.08 ± 0.054	26.04 ± 0.056	26.05 ± 0.054	25.95 ± 0.055
harbara	60	$\frac{22.99 \pm 0.048}{22.02 \pm 0.026}$	22.95 ± 0.091	23.18 ± 0.061	23.80 ± 0.067	23.80 ± 0.071	23.78 ± 0.075	23.81 ± 0.070	23.81 ± 0.070	$\frac{23.79 \pm 0.074}{22.69 \pm 0.017}$	23.80 ± 0.073	23.78 ± 0.075
Darbara	20	32.93 ± 0.026 29.36 ± 0.032	33.50 ± 0.018 29.83 ± 0.029	33.66 ± 0.020 29.96 ± 0.032	33.70 ± 0.021 30.23 ± 0.030	33.66 ± 0.022 30.27 ± 0.028	33.53 ± 0.020 30.04 ± 0.029	33.72 ± 0.017 30.27 ± 0.029	$33./4 \pm 0.01/$ 30.27 ± 0.028	33.69 ± 0.017 30.19 ± 0.026	33.66 ± 0.017 30.24 ± 0.027	33.44 ± 0.017 30.00 ± 0.030
	40	25.68 ± 0.047	25.78 ± 0.029 25.78 ± 0.048	25.79 ± 0.047	26.46 ± 0.043	26.51 ± 0.040	26.51 ± 0.039	30.27 ± 0.029 26.52 ± 0.040	26.52 ± 0.040	26.51 ± 0.039	26.51 ± 0.040	26.51 ± 0.039
	60	23.50 ± 0.032	23.17 ± 0.039	23.57 ± 0.034	24.13 ± 0.037	24.15 ± 0.035	24.16 ± 0.035	24.15 ± 0.036	24.15 ± 0.036	24.16 ± 0.035	24.16 ± 0.035	24.16 ± 0.035
boat	10	31.78 ± 0.015	32.73 ± 0.019	32.81 ± 0.018	32.82 ± 0.017	32.72 ± 0.015	32.61 ± 0.016	32.73 ± 0.018	32.75 ± 0.018	32.72 ± 0.017	32.65 ± 0.017	32.49 ± 0.017
	20	28.40 ± 0.017	29.14 ± 0.017	29.23 ± 0.019	29.37 ± 0.015	29.34 ± 0.015	29.05 ± 0.015	29.30 ± 0.018	29.30 ± 0.018	29.25 ± 0.017	29.27 ± 0.017	28.95 ± 0.018
	40	21.95 ± 0.053 23.64 ± 0.025	25.45 ± 0.021 23.11 ± 0.028	25.45 ± 0.021 23.72 ± 0.026	26.01 ± 0.016 24.01 ± 0.026	25.99 ± 0.016 24.01 ± 0.025	25.92 ± 0.014 23.99 + 0.025	25.98 ± 0.012 24.01 ± 0.025	25.98 ± 0.012 24.01 + 0.025	25.95 ± 0.013 24.00 ± 0.025	25.96 ± 0.012 24.00 ± 0.025	25.92 ± 0.014 23.99 ± 0.025
hill	10	$\frac{23.04 \pm 0.023}{31.87 \pm 0.029}$	$\frac{23.11 \pm 0.028}{32.63 \pm 0.020}$	$\frac{23.72 \pm 0.020}{32.67 \pm 0.019}$	32.71 ± 0.018	32.61 ± 0.023	32.47 ± 0.014	$\frac{24.01 \pm 0.023}{32.67 \pm 0.016}$	32.67 ± 0.016	32.64 ± 0.016	$\frac{24.00 \pm 0.025}{32.55 \pm 0.015}$	32.34 ± 0.013
	20	28.82 ± 0.022	29.23 ± 0.031	29.23 ± 0.031	$\textbf{29.48} \pm \textbf{0.024}$	$\textbf{29.48} \pm \textbf{0.023}$	29.29 ± 0.023	29.45 ± 0.021	29.45 ± 0.021	29.41 ± 0.023	29.42 ± 0.022	29.25 ± 0.024
	40	25.91 ± 0.022	25.70 ± 0.026	25.98 ± 0.024	26.36 ± 0.024	$\textbf{26.38} \pm \textbf{0.022}$	26.37 ± 0.022	26.38 ± 0.022	$\textbf{26.38} \pm \textbf{0.022}$	26.37 ± 0.022	26.38 ± 0.022	26.37 ± 0.022
counta	60	$\frac{24.25 \pm 0.017}{21.80 \pm 0.014}$	$\frac{23.45 \pm 0.013}{22.76 \pm 0.000}$	24.32 ± 0.018	24.60 ± 0.024	24.59 ± 0.024	24.60 ± 0.024	24.59 ± 0.022	24.59 ± 0.022	24.60 ± 0.024	24.60 ± 0.024	24.60 ± 0.024
coupie	20	31.80 ± 0.014 28.14 ± 0.023	32.76 ± 0.009 28.93 ± 0.023	32.81 ± 0.009 28.93 ± 0.023	32.85 ± 0.010 29 16 ± 0.028	32.80 ± 0.011 29 16 ± 0.029	32.71 ± 0.011 28.86 ± 0.030	32.76 ± 0.013 29.11 ± 0.028	32.77 ± 0.013 29.11 ± 0.028	32.75 ± 0.012 29.07 ± 0.029	32.72 ± 0.011 29.08 ± 0.029	32.39 ± 0.012 28 76 ± 0.030
	40	24.93 ± 0.035	25.03 ± 0.025 25.03 ± 0.026	25.05 ± 0.025 25.05 ± 0.026	25.49 ± 0.026	25.50 ± 0.020	25.44 ± 0.032	25.48 ± 0.028	25.48 ± 0.028	25.47 ± 0.029 25.47 ± 0.030	25.47 ± 0.030	25.43 ± 0.030
	60	23.25 ± 0.037	22.76 ± 0.043	23.29 ± 0.036	23.59 ± 0.045	$\textbf{23.60} \pm \textbf{0.044}$	23.59 ± 0.044	23.60 ± 0.044	$\textbf{23.60} \pm \textbf{0.044}$	23.59 ± 0.044	$\textbf{23.60} \pm \textbf{0.044}$	23.59 ± 0.044
fingerprint	10	30.27 ± 0.017	30.87 ± 0.015	30.87 ± 0.016	30.84 ± 0.016	30.80 ± 0.017	30.57 ± 0.016	30.88 ± 0.018	30.88 ± 0.019	30.83 ± 0.019	30.81 ± 0.020	30.50 ± 0.018
	20	26.64 ± 0.010	27.06 ± 0.014	27.06 ± 0.014	27.04 ± 0.014	27.12 ± 0.012	26.72 ± 0.013	27.10 ± 0.012 24.05 ± 0.022	27.10 ± 0.012 24.05 ± 0.022	26.93 ± 0.013	27.05 ± 0.013	26.70 ± 0.012
	40	23.20 ± 0.018 20.93 ± 0.034	23.08 ± 0.024 21.44 ± 0.029	23.08 ± 0.024 21.44 + 0.029	23.96 ± 0.023 21.85 ± 0.041	24.00 ± 0.022 21.97 ± 0.037	24.03 ± 0.023 21.98 + 0.037	24.03 ± 0.022 21.98 + 0.037	24.03 ± 0.022 21.98 ± 0.037	24.03 ± 0.022 21.98 + 0.037	24.03 ± 0.022 21.98 ± 0.037	24.03 ± 0.023 21.98 ± 0.037
MRI	10	40.06 ± 0.043	39.19 ± 0.040	40.81 ± 0.033	40.89 ± 0.032	40.83 ± 0.034	40.71 ± 0.029	40.79 ± 0.040	40.85 ± 0.038	40.83 ± 0.037	40.81 ± 0.038	40.61 ± 0.032
	20	36.14 ± 0.067	34.47 ± 0.047	36.57 ± 0.062	36.70 ± 0.063	36.74 ± 0.065	36.60 ± 0.068	36.74 ± 0.063	$\textbf{36.77} \pm \textbf{0.064}$	36.64 ± 0.068	36.72 ± 0.066	36.59 ± 0.067
	40	32.22 ± 0.067	29.33 ± 0.055	32.31 ± 0.071	32.53 ± 0.069	32.53 ± 0.069	32.52 ± 0.069	32.49 ± 0.064	32.54 ± 0.070	32.52 ± 0.069	32.54 ± 0.069	32.52 ± 0.069
abbey	00 10	$\frac{29.57 \pm 0.056}{29.31 \pm 0.035}$	26.13 ± 0.058	29.68 ± 0.057	29.88 ± 0.056 29.86 ± 0.031	29.88 ± 0.056 29.87 ± 0.033	29.88 ± 0.056 29.38 ± 0.030	$\frac{29.76 \pm 0.060}{29.92 \pm 0.028}$	29.86 ± 0.059 29.92 ± 0.028	29.88 ± 0.056 29.83 ± 0.032	29.88 ± 0.056 29.83 ± 0.033	29.88 ± 0.056 29.34 ± 0.030
ubbey	20	25.53 ± 0.034	25.91 ± 0.029	25.90 ± 0.029 25.91 ± 0.036	25.87 ± 0.031	25.80 ± 0.034	25.27 ± 0.030	25.87 ± 0.028	25.92 ± 0.028 25.87 ± 0.034	25.72 ± 0.032	25.71 ± 0.034	25.16 ± 0.033
	40	22.85 ± 0.036	22.94 ± 0.031	22.94 ± 0.031	23.10 ± 0.030	23.14 ± 0.024	23.10 ± 0.024	23.13 ± 0.025	23.13 ± 0.025	23.12 ± 0.025	23.12 ± 0.025	23.10 ± 0.025
	60	21.60 ± 0.034	$\underline{21.30\pm0.031}$	$\underline{21.58 \pm 0.032}$	21.83 ± 0.038	21.85 ± 0.039	21.85 ± 0.039	21.85 ± 0.039	21.85 ± 0.039	21.85 ± 0.039	21.85 ± 0.039	21.85 ± 0.039
airplane	10	31.42 ± 0.077	32.47 ± 0.088	32.54 ± 0.086	32.49 ± 0.085	32.50 ± 0.084	32.05 ± 0.082	32.36 ± 0.078	32.39 ± 0.077	32.30 ± 0.085	32.33 ± 0.084	31.78 ± 0.082
cuoin	20 40	$2/.52 \pm 0.10/$ 24.62 ± 0.112	28.53 ± 0.110 24.86 ± 0.124	28.60 ± 0.117 24.91 ± 0.105	28.70 ± 0.130 25.26 ± 0.113	28.58 ± 0.129 25 28 + 0.112	28.21 ± 0.119 25.24 ± 0.121	28.57 ± 0.133 25.25 ± 0.120	28.57 ± 0.133 25.25 ± 0.120	28.45 ± 0.127 25.25 ± 0.118	28.46 ± 0.127 25.25 ± 0.117	28.11 ± 0.121 25.24 ± 0.120
	60	22.89 ± 0.185	24.60 ± 0.124 22.60 ± 0.176	22.97 ± 0.194	23.36 ± 0.206	23.39 ± 0.192	23.38 ± 0.193	23.38 ± 0.120 23.38 ± 0.193	23.38 ± 0.193	23.38 ± 0.193	23.38 ± 0.193	23.24 ± 0.120 23.38 ± 0.193
airport	10	32.79 ± 0.042	33.41 ± 0.038	33.58 ± 0.036	33.51 ± 0.038	33.58 ± 0.043	33.09 ± 0.045	33.26 ± 0.041	33.34 ± 0.041	33.23 ± 0.044	33.37 ± 0.044	32.86 ± 0.040
terminal	20	28.74 ± 0.071	29.58 ± 0.050	29.72 ± 0.044	29.95 ± 0.055	29.97 ± 0.054	29.67 ± 0.049	29.70 ± 0.054	29.70 ± 0.054	29.69 ± 0.058	29.74 ± 0.057	29.42 ± 0.047
	40	24.78 ± 0.060	25.21 ± 0.055	25.21 ± 0.055	25.70 ± 0.077	25.71 ± 0.077	25.55 ± 0.070	25.61 ± 0.074	25.61 ± 0.074	25.59 ± 0.074	25.60 ± 0.074	25.51 ± 0.076
allev	10	$\frac{22.72 \pm 0.061}{37.58 \pm 0.075}$	$\frac{22.59 \pm 0.049}{37.65 \pm 0.053}$	22.80 ± 0.063 37.90 ± 0.070	$\frac{23.22 \pm 0.072}{38.44 \pm 0.077}$	$\frac{23.23 \pm 0.073}{38.48 \pm 0.076}$	$\frac{23.20 \pm 0.070}{38.42 \pm 0.082}$	$\frac{23.21 \pm 0.070}{38.44 \pm 0.079}$	$\frac{23.21 \pm 0.070}{38.44 \pm 0.079}$	$\frac{23.21 \pm 0.071}{38.44 \pm 0.079}$	$\frac{23.21 \pm 0.071}{38.45 \pm 0.078}$	$\frac{23.20 \pm 0.070}{38.41 \pm 0.081}$
	20	33.90 ± 0.087	33.35 ± 0.087	34.15 ± 0.102	34.59 ± 0.101	34.60 ± 0.101	34.60 ± 0.101	34.59 ± 0.102				
	40	31.06 ± 0.056	28.67 ± 0.123	31.12 ± 0.057	31.27 ± 0.058	31.26 ± 0.059	31.26 ± 0.060	31.16 ± 0.077	31.22 ± 0.064	31.26 ± 0.059	31.27 ± 0.060	31.26 ± 0.060
	60	29.65 ± 0.101	25.84 ± 0.110	29.68 ± 0.101	$\underline{29.77 \pm 0.102}$	$\underline{29.76 \pm 0.102}$	$\underline{29.77 \pm 0.102}$	29.46 ± 0.116	$\underline{29.72\pm0.102}$	$\underline{29.77 \pm 0.102}$	$\underline{29.77 \pm 0.103}$	$\underline{29.77 \pm 0.102}$
amphitheater	10	32.27 ± 0.038	32.87 ± 0.038	33.02 ± 0.034	32.94 ± 0.035	32.90 ± 0.036	32.39 ± 0.050	32.94 ± 0.040	32.96 ± 0.037	32.83 ± 0.052	32.81 ± 0.049	32.23 ± 0.069
	20 40	28.70 ± 0.089 25.88 ± 0.135	26.94 ± 0.061 25.40 ± 0.093	29.09 ± 0.064 25.90 ± 0.126	29.12 ± 0.071 26.02 ± 0.157	29.02 ± 0.079 26.02 ± 0.173	28.38 ± 0.078 25.99 ± 0.175	29.11 ± 0.073 26.06 ± 0.158	29.11 ± 0.073 26.06 ± 0.158	28.94 ± 0.081 26.01 ± 0.173	28.91 ± 0.079 26.01 ± 0.172	28.52 ± 0.077 25.99 ± 0.175
	60	24.62 ± 0.077	23.45 ± 0.099	23.50 ± 0.120 24.64 ± 0.080	24.72 ± 0.081	24.72 ± 0.082	23.77 ± 0.173 24.72 ± 0.082	24.72 ± 0.096	24.73 ± 0.003	24.72 ± 0.082	24.72 ± 0.082	23.77 ± 0.173 24.72 ± 0.082

C. Performance Studies With Bias-Variance Trade-Off

The bias-variance trade-off was investigated using many natural images. As shown in Fig. 1, a neighborhood size B was used to estimate p_i using the LJS method [33], and this was a significant factor when determining the bias. This tendency was also observed for the other different natural images, as illustrated in Fig. 4. Increasing B in the LJS method moved the bias-variance trade-off curves in the bottom right direction, meaning that the bias increased and the variance decreased. However, the role of the smoothing parameter h changed in the LJS method. Unlike in classical NLM method (see the NLM bias-variance curve in Fig. 1), increasing the smoothing parameter h beyond a certain point in the LJS method did not further decrease the variance in any of the natural images that

we tested. This is because increasing h will also increases the p_i values so that the resulting LJS estimator becomes closer to the noisy input image y_i due to the lack of bounds for the self-weights.

Our proposed methods (LMM-DB, LMM-RP) yielded trade-off curves that have decreased variances for increasing values of the smoothing parameter h. Figure 5 shows the trade-off curves for the cameraman example for different methods (LMM-DB, LMM-RP), different neighborhood sizes (B = 2, 7), and different noise levels ($\sigma = 10, 40$). Our proposed methods yielded bias-variance curves that were less than or equal to those in the LJS method for fixed B and σ . This tendency was also observed with other natural images, as illustrated in Fig. 4. It was important to choose appropriate



Fig. 4. Bias-variance curves for natural images using LJS₊ [33] and our proposed LMM – DB^{one} and LMM – RP^{one} methods with a noise level of $\sigma = 10$. (a) *couple*. (b) *montage*. (c) *lena*. (d) *pepper*. (e) *house*. (f) *MRI*.

neighborhood sizes B in order for the LJS method to obtain a certain level of bias, but our proposed methods were able to achieve that same level of bias by adjusting the smoothing parameter h, which was the same as in classical NLM. Based on our results, it appears that the use of LMM-RP has slightly more advantages than using LMM-DB in terms of the PSNR, as shown in Table I, and the bias-variance trade-off curves, as shown in Fig. 5, for high noise levels.

D. Performance Studies With Visual Quality Assessment

The most important improvements in our proposed LMM-DB and LMM-RP methods when compared to the LJS method were achieved in terms of the visual quality. Figure 6 (a) shows the true cameraman image (left) and the noisy image (right) with a noise level of $\sigma = 10$. Figure 6 (b) presents the filtered images using the LJS method [33] with B = 2 and B = 7. Severe artifacts were observed in the background areas when using B = 2, and these artifacts were reduced when using B = 7. However, there were still some artifacts near the edges of objects. Our proposed LMM-DB and LMM-RP methods exhibited fewer image artifacts than were observed in the images processed using the LJS method for both B = 2, 7. This tendency was observed in many of the natural images, as shown in Fig. 7, especially in the high intensity flat areas. PSNR improvements in the LJS method were achieved with severe (when B = 2) or mild (when B = 7) artifacts; however, our proposed methods achieved both a high



Fig. 5. Bias-variance curves for LMM – DB^{one} and LMM – RP^{one} for comparison with LJS₊ for two neighborhood sizes B = 2, 7 and two noise levels $\sigma = 10, 40$. (a) LMM – DB^{one}($\sigma = 10$). (b) LMM – DB^{one}($\sigma = 40$). (c) LMM – RP^{one}($\sigma = 10$). (d) LMM – RP^{one}($\sigma = 40$).

PSNR and significantly reduced visual artifacts. This ability to reduce the number of visual artifacts in a denoised image is important in some applications, such as diagnostic medical imaging.



(d)

Fig. 6. True, noisy ($\sigma = 10$), and filtered images using LJS₊ [33], and the proposed LMM – DB^{one} and LMM – RP^{one}. (a) True and noisy images ($\sigma = 10$). (b) LJS₊ [33]. (c) Proposed LMM – DB^{one}. (d) Proposed LMM – RP^{one}.

E. Maximum Self-Weights: One vs. Stein's

Two maximum self-weights were proposed for use: the value one in (14) that was proposed in [4], and Stein's in (15) that was proposed in [32]. Figure 8 shows that the LMM – DB^{one} method yielded an improved bias-variance curve and PSNR than did the LMM – DB^{stein} method when the noise levels were low. For high noise levels $\sigma = 40$, the LMM – DB^{stein} method yielded an improved PSNR and bias-variance curve than did the LMM – DB^{one} method. However, these differences were not significant, as also illustrated in terms of the PSNR in Table I. In terms of the visual quality, no significant differences were observed between the two methods.

TABLE IIPERCENTAGE (%) OF $c(||\mathbf{s}||)$ THAT EXCEED 2 USING LMM – DBAND LMM – RP METHODS, $\sigma = 10, B = 2$

		$\sigma = 10$	$\sigma = 20$	$\sigma = 40$	$\sigma = 60$
	LMM-DB one	0.32	0.04	0.03	0.05
cameraman	LMM-DB ^{stein}	0.85	0.66	0.21	0.18
Guardina	LMM-DB ^{one}	0.00	0.00	0.00	0.00
Jingerprini	LMM-DB ^{stein}	0.30	0.13	0.09	0.02
MDI	LMM-DB one	0.10	0.05	0.10	0.13
MKI	LMM-DB ^{stein}	0.18	0.16	0.16	0.16
		$\sigma = 10$	$\sigma = 20$	$\sigma = 40$	$\sigma = 60$
	LMM-RP ^{one}	$\frac{\sigma = 10}{0.25}$	$\frac{\sigma = 20}{0.04}$	$\frac{\sigma = 40}{0.01}$	$\frac{\sigma = 60}{0.00}$
cameraman	LMM-RP ^{one} LMM-RP ^{stein}	$\sigma = 10$ 0.25 1.07	$\sigma = 20$ 0.04 0.90	$\sigma = 40$ 0.01 0.20	$\sigma = 60$ 0.00 0.22
cameraman	LMM-RP ^{one} LMM-RP ^{stein} LMM-RP ^{one}	$\sigma = 10$ 0.25 1.07 0.01	$\sigma = 20$ 0.04 0.90 0.00	$\sigma = 40$ 0.01 0.20 0.00	$\sigma = 60$ 0.00 0.22 0.00
cameraman fingerprint	LMM-RP ^{one} LMM-RP ^{stein} LMM-RP ^{one} LMM-RP ^{stein}	$\sigma = 10$ 0.25 1.07 0.01 0.27	$\sigma = 20$ 0.04 0.90 0.00 0.19	$\sigma = 40$ 0.01 0.20 0.00 0.13	$\sigma = 60$ 0.00 0.22 0.00 0.03
cameraman fingerprint	LMM-RP ^{one} LMM-RP ^{stein} LMM-RP ^{one} LMM-RP ^{stein} LMM-RP ^{one}	$ \begin{aligned} \sigma &= 10 \\ 0.25 \\ 1.07 \\ 0.01 \\ 0.27 \\ 0.09 \\ \end{aligned} $	$\sigma = 20$ 0.04 0.90 0.00 0.19 0.05	$\sigma = 40$ 0.01 0.20 0.00 0.13 0.07	$\sigma = 60$ 0.00 0.22 0.00 0.03 0.09

F. "Practical" Minimax Estimator

The proposed LMM-DB and LMM-RP methods are minimax with respect to $||\mathbf{s}|| \le Y_4$ and $||\mathbf{s}|| \le Y_3$, respectively, as shown in Fig. 3. However, these conditions impose upper bounds for the smoothing parameters h and the optimal h^* , which means that the smoothing parameter values that yield the best PSNR may not be achievable. We empirically investigated this issue using many natural images.

Table II shows the ratio (percentage unit) of the number of pixels for which $c(||\mathbf{s}||) > 2$ to the total number of pixels in the cameraman, fingerprint, and MRI images when the optimal h^* for the highest PSNR was chosen based on the true images for the proposed LMM-DB and LMM-RP methods. For most of the pixels, the LMM-DB and LMM-RP values were minimax. The relationship between the percentage of pixels with $c(||\mathbf{s}||) > 2$ and the root mean squared error (RMSE) is illustrated in Fig. 9 for the cameraman and MRI images. Surprisingly, the optimal global smoothing parameters h for the lowest RMSE point (or the highest PSNR) of the LMM-DB and LMM-RP methods are very close to the smoothing parameters h such that the percentage of $c(||\mathbf{s}||) > 2$ is 0.1%. This phenomenon was not only observed in these two images. As shown in Table III, the pixel percentage of $c(||\mathbf{s}||) > 2$ that do not require knowledge of the true image can still determine smoothing parameters that are able to yield comparable PSNR values to the best PSNR values obtained by using the optimal smoothing parameters calculated based on knowledge of the true image. This was observed in all of the natural images used in our simulations, with different noise levels, and when B = 2 was used. However, the criteria of using the pixel percentage of $c(||\mathbf{s}||) > 2$ did not work very well for B = 7 in our simulations. These criteria can be potentially used when choosing a global smoothing parameter with our proposed methods as a heuristic approach without knowing the true image.

G. Computation Time for Algorithms

Table IV reports the computation time of the proposed methods in comparison with the classical NLM and LJS_+ .





Fig. 7. Filtered results using LJS₊ [33] and the proposed LJS – RP^{one} method with a noise level of $\sigma = 10$ and neighborhood size B = 2. (a) *couple*. (b) *montage*. (c) *lena*. (d) *pepper*. (e) *house*. (f) *MRI*.



Fig. 8. Bias-variance curves and PSNR vs. varying neighborhood sizes (*B*) using classical NLM (only in the PSNR figure), LJS, and the proposed LMM – DB^{stein} vs. LMM – DB^{one} for the cameraman example.

We used 8 threads (Intel Core i7 2.8 GHz) when computing the patch distances for all methods. The local block size was B = 2, the patch size was 7×7 , and the window size was 31×31 . All parameters were fixed for all of results presented in this section. Adjusting these parameters can greatly reduce the running time. For example, setting B = 4, the patch size to 5×5 , and the window size to 13×13 reduces the computation time of the proposed methods to 0.60, 1.12, and 2.91 seconds (s) for 128^2 , 256^2 , and 512^2 images, respectively. However, analytically, the classical NLM requires $3|\mathbf{P}||\Omega| +$ $4|\Omega| - 1$ operations per pixel where $|\Omega|$ is the number of elements in Ω_i and LJS₊ requires $3|\mathbf{P}||\Omega| + 4|\Omega| + 3|\mathbf{B}| + 5$ operations per pixel. It is reported in [33] that the additional operations for LJS_+ (3|**B**| + 6 operations) were negligible compared to the NLM filtering computation $(3|\mathbf{P}||\Omega|+4|\Omega|-$ 1 operations). Analytically, the additional computation for LMM – DB and LMM – RP is $3|\mathbf{B}| + 7$ operations, which is almost the same as the additional computation for LJS₊. Therefore, further implementation optimization is possible by exploiting the redundant computation of the patch distances for the minimax estimator and NLM weights.

VI. DISCUSSION

The classical NLM method was a significant work in image denoising [4], and required the determination of two important parameters for good denoising performance: a smoothing parameter and a self-weight value. The LJS method proposed

TABLE III

The PSNR Values (dB) of the Proposed Methods With B = 2 When Choosing the Smoothing Parameter so as to Yield the Highest PSNR Using the True Image (TRUE), and When Choosing the Smoothing Parameter so as to Yield the Percentage of $c(\|s\|) > 2$ to be 0.1% (ESTIMATED) for Different Noise Levels

		LMM-DB ^{one}		LMM-RP one			
	σ	TRUE	ESTIMATED	TRUE	ESTIMATED		
cameraman	10	33.35	33.32	33.22	33.30		
	20	29.47	29.47	29.30	29.45		
	40	25.91	25.90	26.16	26.01		
	60	23.43	23.43	23.78	23.60		
lena	10	34.72	34.74	34.73	34.78		
	20	31.22	31.20	31.47	31.37		
	40	27.74	27.74	28.07	27.88		
	60	25.60	25.63	25.07	25.82		
montage	10	35.55	35.56	35.51	35.34		
	20	31.24	31.20	31 33	31.32		
	10	26.00	26.08	27.26	27.12		
	40	20.99	20.98	27.20	27.15		
house	10	25.22	25.37	25.20	24.10		
nouse	20	33.32	33.33	33.30	33.37		
	20	32.00	27.70	22.30	28.06		
	40	27.62	27.79	26.33	28.00		
nenners	10	23.25	23.25	23.13	23.57		
peppers	10	33.30	33.34	33.48	33.39		
	20	29.81	29.80	29.91	29.95		
	40	25./1	25.67	20.10	23.89		
harborn	00	23.15	23.02	23.81	23.24		
Darbara	10	33.67	33.68	33.62	33.73		
	20	29.94	29.82	30.23	30.03		
	40	25.79	25.69	26.47	25.92		
	60	23.56	23.48	24.14	23.67		
boat	10	32.80	32.80	32.73	32.79		
	20	29.22	29.18	29.36	29.22		
	40	25.44	25.60	25.99	25.75		
	60	23.76	23.76	24.05	23.91		
hill	10	32.66	32.64	32.60	32.61		
	20	29.24	29.17	29.49	29.30		
	40	26.01	25.99	26.38	26.05		
	60	24.35	24.35	24.63	24.45		
couple	10	32.82	32.80	32.79	32.81		
	20	28.89	28.70	29.12	28.85		
	40	25.08	25.05	25.52	25.10		
	60	23.26	23.19	23.56	23.28		
fingerprint	10	30.86	30.84	30.66	30.80		
	20	27.07	26.86	27.05	26.96		
	40	23.69	23.24	23.96	23.38		
	60	21.47	20.77	21.92	21.02		
MRI	10	40.83	40.83	40.77	40.90		
	20	36.59	36.60	36.71	36.74		
	40	32.36	32.36	32.58	32.56		
	60	29.64	29.63	29.83	29.80		
abbey	10	30.01	29.98	29.98	29.94		
	20	25.89	25.74	25.91	25.60		
	40	22.96	22.80	23.03	22.81		
	60	21.58	21.58	21.71	21.65		
airplane	10	32.47	32.45	32.50	32.49		
cabin	20	28.51	28.35	28.61	28.45		
	40	24.74	24.68	25.00	24.81		
	60	23.13	23.04	23.34	23.30		
airport	10	33.64	33.64	33.68	33.68		
terminal	20	29.63	29.53	29.83	29.63		
	40	25.19	25.04	25.51	25.07		
	60	22.95	22.94	23 23	23.08		
allev	10	37.94	37.87	38.28	38.06		
	20	34.15	34.10	34 44	34 25		
	20 20	31.06	31.03	31 10	31.19		
	60	20 53	29.10	20 50	20.40		
amphitheater	10	33.00	32.07	33.00	32.04		
r	20	20.19	20.08	20 31	20.13		
	20 10	27.10	25.00	29.51	25.15		
	40	25.90	23.90	23.98	23.93		

by Wu *et al.* [33] developed a state-of-the-art method for selfweight determination using JS estimation [37] and yielded superior results in terms of the PSNR compared to the other existing methods. However, since the LJS method did not impose an upper bound for self-weight estimation, the bias could no longer be controlled by the smoothing parameter, which resulted in visual quality degradation. Our proposed methods based on the Baranchik's minimax theorem [34] yielded comparable PSNR results to the state-of-the-art LJS method. By imposing upper bounds for the self-weights,



Fig. 9. Comparison plots of the RMSE vs. the smoothing parameter h and the percentage of c(||s||) > 2 vs. the same smoothing parameter when using LMM-DB and LMM-RP with B = 2 and $\sigma = 10$. (a) cameraman LMM – DB^{one}. (b) MRI LMM – DB^{one}. (c) cameraman LMM – RP^{one}. (d) MRI LMM – RP^{one}.

TABLE IV Execution time (s) Comparison. This Will Vary With Parameter Selection

Image size	Classical NLM	LJS_{+}	LMM-DB one	LMM-RP one
128*128	0.65	0.89	0.90	0.90
256*256	1.57	2.37	2.39	2.38
512*512	4.90	7.14	7.19	7.18

the bias-variance trade-off was able to be controlled by a smoothing parameter, and substantial visual artifact reduction was achieved.

The focus of this article was self-weight parameter selection in the classical NLM filter with theoretical justification. As discussed in the Introduction, there are other factors that affect the performance of NLM based filters, and we expect that our proposed methods would not be able to achieve state-of-the-art denoising performance if there were no other optimizations performed except the self-weights. Indeed, our proposed methods with one patch size (non-adaptive neighborhood) and one global smoothing parameter were not able to achieve the level of denoising performance of the state-of-theart denoising methods such as BM3D [6]. However, when our proposed methods have incorporated some of the other factors into the NLM filters, such as local smoothing parameters and adaptive neighborhoods [10], they have great potential to achieve significantly improved denoising performance.

The minimax property of our proposed methods depends on the choice of smoothing parameters. When using sufficiently small smoothing parameters, the LMM-DB and LMM-RP methods are "practically" minimax according to Baranchik's theorem [34]. However, when large smoothing parameters are used, there may be some pixels that are not minimax for selfweight estimation. More empirical investigation showed that the optimal global smoothing parameter h that yielded the best PSNR only resulted in a very small portion of the pixels that did not have minimax self-weight estimators. In fact, this can be used as a useful heuristic when choosing a good smoothing parameter since testing the minimax properties of our proposed methods does not require the true image. More theoretical analysis for this observation, or a statistical analysis using many natural images as shown in [41], are potential extensions of this work. Therefore, our proposed methods do not only provide an optimal way to determine self-weights, but also provide a heuristic way to determine a good smoothing parameter.

VII. CONCLUSION

We proposed two methods, LMM-DB, LMM-RP, to determine the self-weights of NLM filters that are "practically" minimax, and this methods yielded a comparable PSNR, better bias-variance trade-offs, and reduced visual quality artifacts when compared to the results obtained using the state-of-theart LJS method. Our methods also provide a potentially useful heuristic way to determine a global smoothing parameter without knowledge of the original image.

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