

000 LATENT G-COMPUTATION FOR POTENTIAL OUT- 001 COMES DISTRIBUTIONAL ESTIMATION UNDER TIME- 002 VARYING TREATMENTS 003

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ABSTRACT

012
 013 Estimating individualized potential outcomes (POs) under time-varying treat-
 014 ments is central to fields like medicine, marketing, and public policy, where de-
 015 cisions must account for uncertainty rather than just point forecasts. We intro-
 016 duce a latent g-computation estimator for discrete-time, individualized PO dis-
 017 tributions. Under standard longitudinal identification assumptions and a latent
 018 factorization/context-sufficiency condition—essentially the usual expressivity as-
 019 sumption for conditional VAEs—, we show that a rollout entirely in latent space
 020 targets the same interventional distribution as the classical g-formula, while never
 021 autoregressing covariates in data space. We further derive a total-variation er-
 022 ror-propagation bound proving that, for a given one-step approximation error,
 023 latent rollouts exhibit more favorable long-horizon behavior than data-space au-
 024 toregressive g-computation. We instantiate this estimator as G-Latent, which re-
 025 places G-Net’s residual pools (Li et al., 2021) with a conditional VAE that learns
 026 history- and treatment-conditioned outcome distributions at each time. To enhance
 027 expressivity, we adapt an infinite-mixture asymmetric Laplace (ALD) parameter-
 028 ization (An & Jeon, 2023) to the time-series setting, and we decouple sequence
 029 encoding (a transformer over the observed history) from a lightweight GRU latent
 030 rollout with selective decoding, enabling fast Monte Carlo sampling over multiple
 031 horizons. We evaluate G-Latent in semi-synthetic and real-world datasets, finding
 032 that it yields better calibrated and more accurate predictive PO distributions than
 033 strong baselines, while reducing inference-time cost.
 034

1 INTRODUCTION

035
 036 Estimating individualized potential outcomes under time-varying treatments is central to data-rich
 037 domains such as precision medicine, marketing, education, and public policy, where longitudinal
 038 records capture detailed sequences of covariates, interventions, and responses. While recent neu-
 039 ral approaches address time-dependent confounding and long-range dependencies, most return only
 040 point estimates—typically conditional means (Melnychuk et al., 2022; Bouchattaoui et al., 2023)—or
 041 consider only *epistemic* (model) uncertainty. Modeling epistemic uncertainty is valuable for flagging
 042 low-confidence regions or detecting out-of-distribution inputs; however, it leaves *aleatoric* (data)
 043 uncertainty unmodeled, so identical expected outcomes may conceal very different variances, skew-
 044 ness, and tail risks. For risk-sensitive decisions—where clinicians care about adverse-event prob-
 045 abilities, marketers about downside exposure, and policymakers about extreme impacts—ignoring
 046 aleatoric uncertainty limits actionable guidance. We therefore advocate moving beyond mean ef-
 047 fects and purely epistemic views to full, coherent distributional estimates of individualized potential
 048 outcomes across time and variables, enabling transparent, risk-aware decision support.

049 We introduce G-Latent, a model for distributional individualized POs under time-varying treatments
 050 that performs g-computation in latent space. The key idea is a latent rollout: during counterfactual
 051 rollouts, we update the temporal representation using VAE latent variables rather than observed co-
 052 variates, and decode only when needed. This avoids data-space autoregression—reducing accumu-
 053 lation error and making g-computation practical with high-dimensional covariates—while enabling
 054 efficient sampling for many treatment sequences and Monte Carlo (MC) draws. G-Latent learns
 055 per-step conditional distributions non-parametrically via a conditional VAE on past representations.

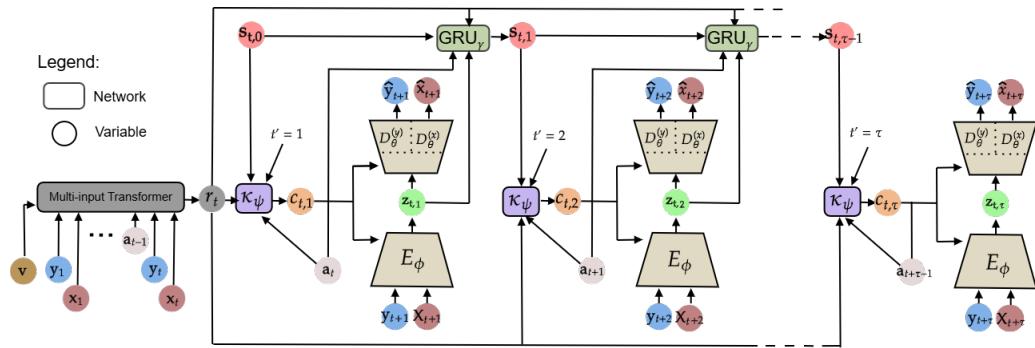


Figure 1: Training-time data flow in G-Latent for a given step t . A multi-input transformer encodes history r_t ; at each step t' within the projection horizon, a context $c_{t,t'}$ feeds a shared conditional VAE. The GRU updates the state using latents $z_{t,t'}$ (not decoded observations). The decoder has outcome (ALD) and covariate (Gaussian) heads.

Following (An & Jeon, 2023) and extending to time series, we parameterize the decoder as an infinite mixture of asymmetric Laplace distributions (ALDs) (Brando et al., 2019), increasing expressivity. In contrast, g-computation baselines such as Li et al. (2021) approximate distributions via mean predictions plus errors from a global residual pool, which can distort individualized distributions, especially under heteroscedasticity. For efficiency, we decouple long-history encoding and short-horizon rollout: a transformer encodes the long prefix once; a lightweight Gated Recurrent Unit (GRU) updates representations across the projection horizon, avoiding repeated transformer passes during sampling. Identifiability follows the g-computation formula under standard assumptions of sequential ignorability, positivity, and consistency.

We summarize our contributions as follows: **1)** We define a novel *latent g-computation estimator* for individualized potential outcome distributions in discrete time under time-varying treatments. Under standard longitudinal identification assumptions and a latent factorization / context-sufficiency condition —essentially the usual expressivity assumption for conditional VAEs— we prove that a rollout entirely in latent space targets the same interventional distribution as the classical g-formula while never autoregressing covariates in data space (Thm. 5.1, Cor. 5.2). To our knowledge, furthermore, ours is the first discrete-time method for individualized distributional POs without global residual pools. **2)** We analyze error propagation for latent vs. data-space implementations of g-computation and derive a total-variation bound showing that, for any fixed one-step approximation error, latent rollouts exhibit more favorable long-horizon behavior than standard autoregressive g-computation (Prop. 5.3), theoretically explaining the improved stability we observe at longer horizons. **3)** We instantiate this estimator as *G-Latent*, a conditional VAE with a transformer history network, a lightweight GRU latent rollout, and an ALD-mixture outcome head adapted from An & Jeon (2023), which together enable flexible individualized outcome distributions and fast Monte Carlo sampling via selective decoding. **4)** We provide an extensive empirical study on semi-synthetic and real-world ICU data, including calibration metrics, runtime comparisons, and an analysis (and correction) of the widely used semi-synthetic MIMIC-III (Melnichuk et al., 2022) benchmark that previously violated positivity. Across datasets, G-Latent improves the quality and calibration of predictive PO distributions relative to strong baselines while reducing inference-time cost.

2 RELATED WORK

Potential outcomes estimation in static settings. In the static setting, there are several methods for individualized PO estimation. Representative modern examples include Yoon et al. (2018); Vansteelandt & Morzywolek (2023); Shalit et al. (2017); Künzel et al. (2019). Although most static PO methods provide only point estimates, some works estimate distributional POs. For instance, papers like Melnichuk et al. (2023); Kennedy et al. (2023) target population-level distributional POs, whereas Ma et al. (2024) learn individualized distributional POs using diffusion models (Yang et al., 2023).

108 **Individualized potential outcomes estimation over time.** Traditionally, causal inference has ad-
 109 dressed time-varying confounders with Marginal Structural Models (MSMs) (Robins et al., 2000),
 110 which rely on inverse probability of treatment weighting (IPTW) (Chesnaye et al., 2022), or G-
 111 computation (Taubman et al., 2009). Lim (2018) improve MSMs by employing RNNs in the mod-
 112 eling of outcomes and propensities. Counterfactual Recurrent Network (CRN) (Bica et al.) in-
 113 corporates adversarial domain training to establish a treatment-invariant representation space using
 114 a gradient reversal layer (Ganin & Lempitsky, 2015). G-Net (Li et al., 2021) combines RNNs
 115 with G-computation to adjust for confounders and estimate dynamic potential outcomes. Causal
 116 Transformer (CT) (Melnychuk et al., 2022) follows the treatment-invariant representation idea from
 117 CRN and incorporates transformers to process time series and a Counterfactual Domain Confusion
 118 (CDC) loss (Tzeng et al., 2015). Other works that also follow this idea are Wang et al. (2024), which
 119 adopts a novel Temporal Integration Predicting strategy and focuses on continuous treatments, and
 120 El Bouchattaoui et al. (2024), which introduces an RNN backbone trained with Contrastive Predic-
 121 tive Coding and an InfoMax objective. Wang et al. (2025) use a state-space architecture (Mamba)
 122 (Gu & Dao, 2024) that employs covariate-based decorrelation toward selective parameters to reduce
 123 confounding bias. Huang et al. (2024) provide an empirical evaluation of balancing strategies. On
 124 the other hand, Xiong et al. (2024) use a similar approach to G-Net but processing data with trans-
 125 formers instead of RNNs, and Deng et al. (2024) add model uncertainty to the same approach. Hess
 126 et al. (2024) propose a pseudo-outcome regression based on g-formula to obtain individualized POs.
 127 Finally, Frauen et al. propose a series of model-agnostic meta-learners for estimating heterogeneous
 128 treatment effects over time.

129 In parallel to the previous works, another line of research has appeared in recent years that
 130 models the effects of treatments in continuous-time with neural Ordinary Differential Equations
 131 (ODEs). De Brouwer et al. (2022) couples neural ODEs with epistemic uncertainty quantification
 132 for continuous-time predictions. Seedat et al. (2022) learn Controlled Differential Equation (CDE)
 133 dynamics robust to irregular sampling. Hess et al. present Bayesian Neural CDE (BNCDE), which
 134 provides posterior predictive distributions over POs. Finally, Hess & Feuerriegel employ a stabilized
 135 continuous-time IPTW formulation to address time-varying confounding.

136 All the previous works, like ours, assume sequential ignorability (Robins & Hernan, 2008). There
 137 is another line of research that tackles violations of this assumption. Among them, papers like
 138 Peng et al.; Bouchattaoui et al. (2023) are worth mentioning as, like this work, they use the latent
 139 representations of VAEs. However, they do it to infer hidden confounders in settings where they
 140 exist. In contrast, our work uses latent representations to adjust for observed confounders following
 141 G-computation. Finally, Wang et al. present another VAE-based approach that aims at selecting best
 142 treatment sequences by modeling the conditional likelihood of achieving target outcomes.

143 **Uncertainty Quantification in potential outcomes estimation over time.** Some of the afore-
 144 mentioned time-varying methods include some form of uncertainty quantification. Within the
 145 continuous-time works, De Brouwer et al. (2022) handles epistemic uncertainty through variational
 146 Bayesian inference . On the other hand, Hess et al. handles both epistemic uncertainty, with
 147 Bayesian posterior distributions, and aleatoric uncertainty, with a Gaussian outcome head. How-
 148 ever, it does not handle time-varying confounding. Very recently, a new paper appeared (Mu et al.,
 149 2025) that employs diffusion models to model distributional potential outcomes with expert models.

150 As for discrete time models for individualized POs, uncertainty quantification has been mostly ig-
 151 nored. Papers like Melnychuk et al. (2022); Bica et al. handle epistemic uncertainty only through
 152 Monte Carlo (MC) dropout. As for aleatoric uncertainty, G-Net (Li et al., 2021) and its transformer
 153 extension (Xiong et al., 2024) are, to the best of our knowledge, the only models that handle it.
 154 Like our model, G-Net builds on g-computation to generate sequential MC samples. However, its
 155 capacity to properly model PO distributions is limited because it only handles homoscedastic data.
 156 Furthermore, it tends to underperform in comparison with other methods due to an error compounding
 157 problem. Deng et al. (2024) enriches (Transformer) G-Net by adding epistemic uncertainty, but
 158 it suffers from the same problems as (Transformer) G-Net. Finally, Wu et al. (2024) combine VAEs
 159 and diffusion models with IPTW to obtain distributional POs, and Shirakawa et al. (2024) couple
 160 a temporal-difference heterogeneous Transformer with longitudinal Targeted Minimum Loss-based,
 161 allowing to estimate POs confidence intervals, but these works handle only population-level POs, so
 162 they do not fit our setting.

162

3 PROBLEM FORMULATION

163

164 For the variables of our setting, uppercase bold letters (e.g., $\mathbf{X}, \mathbf{A}, \mathbf{Y}$) denote random vectors; low-
165 ercase bold (e.g., $\mathbf{x}, \mathbf{a}, \mathbf{y}$) their realizations, and plain letters denote scalars (e.g., x, y). For latent
166 vectors and learnable representation vectors, we use bold lowercase.

167 **Problem Setting.** We adopt the standard setting for estimating counterfactual outcomes over time
168 (Lim, 2018; Bica et al.; Melnychuk et al., 2022; El Bouchattaoui et al., 2024). Let i index patients
169 with trajectories observed at $t = 1, \dots, T^{(i)}$. At each t we observe time-varying covariates $\mathbf{X}_t^{(i)} \in$
170 \mathbb{R}^{d_x} , treatments $\mathbf{A}_t^{(i)}$, and outcomes $\mathbf{Y}_t^{(i)} \in \mathbb{R}^{d_y}$, as well as static covariates $\mathbf{V}^{(i)}$ (e.g., sex, age,
171 risk factors). Unless needed, we omit the patient index (i). We assume i.i.d. observational data $\mathcal{D} =$
172 $\{(\mathbf{x}_{1:T^{(i)}}, \mathbf{a}_{1:T^{(i)}}, \mathbf{y}_{1:T^{(i)}}, \mathbf{v}^{(i)})\}_{i=1}^N$, with $\mathbf{x}_{1:T^{(i)}}^{(i)} = (\mathbf{x}_1^{(i)}, \dots, \mathbf{x}_{T^{(i)}}^{(i)})$ and analogously for \mathbf{a}, \mathbf{y} .

173 **History and calendar.** We use *start-of-interval* indexing: the treatment \mathbf{A}_t precedes the next measurement
174 ($\mathbf{Y}_{t+1}, \mathbf{X}_{t+1}$). Let the history available *before* choosing \mathbf{A}_t be $\bar{\mathbf{H}}_t = \{\bar{\mathbf{X}}_t, \bar{\mathbf{A}}_{t-1}, \bar{\mathbf{Y}}_t, \mathbf{V}\}$
175 with $\bar{\mathbf{X}}_t = (\mathbf{X}_1, \dots, \mathbf{X}_t)$, $\bar{\mathbf{Y}}_t = (\mathbf{Y}_1, \dots, \mathbf{Y}_t)$, and $\bar{\mathbf{A}}_{t-1} = (\mathbf{A}_1, \dots, \mathbf{A}_{t-1})$. For compactness we
176 sometimes group outcomes and covariates as $\mathbf{L}_t = (\mathbf{Y}_t, \mathbf{X}_t) \in \mathbb{R}^{d_L}$.

177 **Targets.** Let $\tau \geq 1$ denote the projection horizon and $\bar{\mathbf{a}}_{t:t+\tau-1} = (\mathbf{a}_t, \dots, \mathbf{a}_{t+\tau-1})$ a given (non-
178 random) treatment intervention. Most previous works in this setting aim to estimate the conditional
179 mean $\mathbb{E}[\mathbf{Y}_{t+\tau} | \bar{\mathbf{a}}_{t:t+\tau-1} | \bar{\mathbf{H}}_t]$. In contrast, we target the *full conditional distribution*, both at a fixed
180 horizon and jointly across horizons:

181
$$p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+\tau} | \bar{\mathbf{h}}_t), \quad p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+1:t+\tau} | \bar{\mathbf{h}}_t). \quad (1)$$

182 **Assumptions.** We build upon the potential outcomes framework (Rubin, 2005) and its extension to
183 time-varying treatments (Robins et al., 2000). We assume (1) consistency, (2) sequential ignorability/
184 exchangeability, and (3) sequential overlap/positivity (see App. A).

185 **Goal.** We design a novel implementation of g-computation that learns flexible per-step conditionals
186 and generates *coherent fast Monte Carlo samples* from $p^{\bar{\mathbf{a}}}(\cdot | \bar{\mathbf{h}}_t)$, enabling distributional individu-
187 alized potential outcomes without data-space autoregression.

188 **The g-Formula.** Under the assumptions previously specified, for any non-random regime $\bar{\mathbf{a}}_{t:t+\tau-1}$,

189
$$p^{\bar{\mathbf{a}}_{t:t+\tau-1}}(\mathbf{y}_{t+1:t+\tau} | \bar{\mathbf{h}}_t) = \int_{\mathbf{x}_{t+1:t+\tau}} \prod_{s=t}^{t+\tau-1} p(\mathbf{y}_{s+1}, \mathbf{x}_{s+1} | \bar{\mathbf{h}}_s, \mathbf{a}_s) d\mathbf{x}_{t+1:t+\tau}, \quad (2)$$

190 where $\bar{\mathbf{h}}_{s+1} := (\bar{\mathbf{h}}_s, \mathbf{a}_s, \mathbf{y}_{s+1}, \mathbf{x}_{s+1})$.

191

4 LATENT G-COMPUTATION

192

193 In this section, we first define a *latent g-computation estimator* that implements discrete-time g-
194 computation entirely in latent space (Section 4.1). Under a latent factorization / context-sufficiency
195 condition, we show in Section 4.2 that this estimator targets the same interventional distribution
196 as the classical g-formula, while never autoregressing covariates in data space. We then analyze
197 its error propagation and finally instantiate it as a neural model, G-Latent, based on a transformer
198 history network, a conditional VAE, and GRU updates in latent space.

199

4.1 THE LATENT G-COMPUTATION ESTIMATOR

200

201 Consider the g-formula (Eq. 2), which expresses the interventional law under a non-random treat-
202 ment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$ as an iterated integral over one-step conditionals

203
$$p^{\star}(\mathbf{y}_{s+1}, \mathbf{x}_{s+1} | \bar{\mathbf{h}}_s, \mathbf{a}_s), \quad s = t, \dots, t + \tau - 1. \quad (3)$$

204 Standard implementations of g-computation approximate these kernels directly in data space and
205 then perform autoregressive rollouts, repeatedly sampling covariates and feeding them back into the
206 model. We instead ask whether g-computation can be implemented entirely in latent space, so that
207 we never autoregress observed covariates while still targeting the same interventional distribution.

216 **Algorithm 1** Latent g-computation estimator (Monte Carlo rollout)

217

218 1: **Input:** history $\bar{\mathbf{h}}_t$, treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$, horizon τ , samples M , scope $\in \{\text{all}, \text{last}\}$

219 2: $\mathbf{r}_t \leftarrow f_\omega(\bar{\mathbf{h}}_t)$

220 3: **for** $m = 1$ **to** M **do** ▷ Monte Carlo paths

221 4: $\mathbf{s}_{t,0} \leftarrow \mathbf{0}$

222 5: **for** $t' = 1$ **to** τ **do**

223 6: $\mathbf{c}_{t,t'} \leftarrow \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$

224 7: $\mathbf{z}_{t,t'}^{(m)} \sim p_0(\cdot)$ ▷ e.g., $\mathcal{N}(\mathbf{0}, \mathbf{I})$

225 8: **if** scope = all **or** $t' = \tau$ **then**

226 9: **decode** $\mathbf{y}_{t+t'}^{(m)}$, $(\mathbf{x}_{t+t'}^{(m)}) \sim p_\theta(\cdot | \mathbf{z}_{t,t'}^{(m)}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$

227 10: $\mathbf{s}_{t,t'} \leftarrow \Gamma_\gamma(\mathbf{z}_{t,t'}^{(m)}, \mathbf{r}_t, \mathbf{a}_{t+t'-1}, t', \mathbf{s}_{t,t'-1})$

228 11: **Return:** $\{\mathbf{y}_{t+1:t+\tau}^{(m)}\}_{m=1}^M$ if scope = all, else $\{\mathbf{y}_{t+\tau}^{(m)}\}_{m=1}^M$

229

230 Fix a time t and a prediction horizon $\tau \geq 1$. Let $\bar{\mathbf{h}}_t$ denote the observed history up to time t and
 231 $\bar{\mathbf{a}}_{t:t+\tau-1}$ a treatment plan applied from t to $t+\tau-1$. Our latent estimator uses four components: (i)
 232 a *history network* f_ω that maps the observed history to an embedding $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$; (ii) a recurrent
 233 *latent state* $\mathbf{s}_{t,t'}$ summarizing the latent trajectory from t to $t+t'$, initialized as $\mathbf{s}_{t,0} = \mathbf{0}$ and updated
 234 as

$$\mathbf{s}_{t,t'} = \Gamma_\gamma(\mathbf{z}_{t,t'}, \mathbf{r}_t, \mathbf{a}_{t+t'-1}, t', \mathbf{s}_{t,t'-1}) \quad t' = 1, \dots, \tau; \quad (4)$$

235 (iii) a *context map* $\mathbf{c}_{t,t'} = \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$, which collects all information needed by the
 236 one-step decoder at step t' ; and (iv) a conditional decoder p_θ with fixed latent prior p_0 defining
 237 one-step kernels

$$p_\theta(\mathbf{l}_{t+t'} | \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}), \quad \mathbf{l}_{t+t'} = (\mathbf{y}_{t+t'}, \mathbf{x}_{t+t'}), \quad \mathbf{z}_{t,t'} \sim p_0(\cdot). \quad (5)$$

238 Given these components, we implement g-computation by *ancestral sampling of full latent paths*.
 239 For each Monte Carlo replicate, we sample a trajectory of latents $\mathbf{z}_{t,1}, \dots, \mathbf{z}_{t,\tau}$ under the treatment
 240 plan, update the latent state forward in time, and decode outcomes (and optionally covariates) at
 241 selected horizons. Crucially, decoded observations are never fed back into the state; all temporal
 242 dependence flows through $(\mathbf{r}_t, \mathbf{s}_{t,t'})$. In our concrete instantiation (Section 4.3), p_θ and p_0 arise
 243 from a conditional VAE over $(\mathbf{Y}_t, \mathbf{X}_t)$.

244 With our estimator, one can decode at any subset $S \subseteq \{1, \dots, \tau\}$ of relative steps. The latent
 245 rollout and state updates are identical in all cases; only decoding is selective. We parameterize
 246 this choice via an argument *scope* that specifies at which relative steps we decode outcomes. In
 247 this work, we consider two options: *scope=all* corresponds to decoding at all $t' = 1, \dots, \tau$,
 248 while *scope=last* corresponds to decoding only at $t' = \tau$. This selective decoding is useful
 249 computationally: when we are only interested in $\mathbf{y}_{t+\tau}$, choosing *scope=last* avoids decoding at
 250 the intermediate $\tau - 1$ steps, reducing the decoder cost from $O(\tau M)$ to $O(M)$ for M Monte Carlo
 251 paths. More generally, decoding at an arbitrary subset S scales the decoder cost linearly in $|S|$ rather
 252 than in τ .

253 Algorithm 1 defines our latent g-computation estimator: given a history $\bar{\mathbf{h}}_t$ and a treatment plan
 254 $\bar{\mathbf{a}}_{t:t+\tau-1}$, it produces Monte Carlo samples from an interventional distribution induced by the one-
 255 step conditionals $p_\theta(\mathbf{l}_{t+t'} | \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$. In Section 4.2, we state conditions under which this
 256 estimator is equivalent to the classical g-formula and analyze its error propagation. In Section 4.3,
 257 we describe how we instantiate $(f_\omega, \kappa_\psi, p_\theta, \Gamma_\gamma)$ as the neural model G-Latent.

261 4.2 THEORETICAL INSIGHTS

262 We now provide theoretical guarantees that the latent g-computation estimator implements the same
 263 interventional law as the traditional data-space g-formula, and compare its error propagation to a
 264 data-space autoregressive rollout. See full proofs and additional discussion in App. E.

265 **Assumption 4.1.** (Latent factorization and context sufficiency). Fix t and $\tau \geq 1$. Let $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$,
 266 let the latent state $\mathbf{s}_{t,t'}$ and context $\mathbf{c}_{t,t'}$ be defined as in Section 4.1, and consider the one-step
 267 conditional over $\mathbf{l}_{t+t'} = (\mathbf{y}_{t+t'}, \mathbf{x}_{t+t'})$. We assume that the true one-step conditional admits a latent
 268 mixture factorization with a fixed prior p_0 :

$$p^*(\mathbf{l}_{t+t'} | \bar{\mathbf{h}}_{t+t'-1}, \mathbf{a}_{t+t'-1}) = \int p_\theta(\mathbf{l}_{t+t'} | \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}) p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'}. \quad (6)$$

270 (See App. E.1 for the formal statement and further discussion.)
 271

272 Intuitively, this says that once we condition on a sufficiently informative context
 273 $\mathbf{c}_{t,t'}(\bar{\mathbf{h}}_{t+t'-1}, \mathbf{a}_{t+t'-1})$, the decoder family $p_\theta(\cdot | \mathbf{z}, \mathbf{c}, \mathbf{a})$ is rich enough to represent the true
 274 one-step conditional as a mixture over a fixed prior p_0 , as in a standard conditional VAE. This is the
 275 standard conditional VAE modeling assumption and not an additional causal assumption.

276 **Theorem 4.2** (Equivalence of latent and data-space g -computation). *Under the identification as-
 277 sumptions (App. A) and the latent factorization in Eq. 6, for any treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$ and history
 278 $\bar{\mathbf{h}}_t$, Algorithm 1 produces i.i.d. MC samples from the interventional laws identified by the g -formula
 279 (Eq. 2):*

$$\begin{aligned}
 \text{(full path)} \quad p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+1:t+\tau} | \bar{\mathbf{h}}_t) &= \int \prod_{t'=1}^{\tau} p_\theta(\mathbf{y}_{t+t'} | \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}(\mathbf{z}_{t,1:t'-1}), \mathbf{a}_{t+t'-1}) \prod_{t'=1}^{\tau} p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'}, \\
 \text{(fixed horizon)} \quad p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+\tau} | \bar{\mathbf{h}}_t) &= \int p_\theta(\mathbf{y}_{t+\tau} | \mathbf{z}_{t,\tau}, \mathbf{c}_{t,\tau}(\mathbf{z}_{t,1:\tau-1}), \mathbf{a}_{t+\tau-1}) \prod_{t'=1}^{\tau} p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'}.
 \end{aligned} \tag{7}$$

286 Proof. App. E.5.

287 **Corollary 4.3** (Selective decoding is coherent). *Decoding only at $t+\tau$ (scope=last) returns
 288 i.i.d. samples from $p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+\tau} | \bar{\mathbf{h}}_t)$; decoding at any subset $S \subseteq \{1, \dots, \tau\}$ returns the corresponding
 289 marginals $\{p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+t'} | \bar{\mathbf{h}}_t)\}_{t' \in S}$. Proof. App. E.5.*

291 **Error propagation: latent vs. data-space g -computation rollouts.** *Takeaway:* the latent rollout
 292 (Alg. 1) does not amplify local one-step errors, whereas data-space autoregressive (AR) rollouts can,
 293 because they repeatedly decode and re-encode observations.

294 In latent g -computation, the learned one-step kernel is the decoder-induced latent mixture at context
 295 $\mathbf{c}_{t,t'}, K_s^e(\cdot | \bar{\mathbf{h}}_s, \mathbf{a}_s) = \int p_\theta(\cdot | \mathbf{z}, \mathbf{c}_{t,t'}, \mathbf{a}_s) p_0(\mathbf{z}) d\mathbf{z}$ with $s = t+t'-1$; $\mathbf{c}_{t,t'}$ is defined in Sec. 4.1
 296 and the state is updated *through latents only*. As a comparator, we use a data-space AR rollout that
 297 decodes each step and re-feeds (or re-encodes), inducing a single-step Lipschitz AR tail operator
 298 with factors $\{1 + \lambda_j\}$. Let $K_s^*(\cdot | \bar{\mathbf{h}}_s, \mathbf{a}_s)$ denote the true one-step conditional and define $\varepsilon_s :=$
 299 $\sup_{\bar{\mathbf{h}}_s, \mathbf{a}_s} \text{TV}(K_s^*(\cdot | \bar{\mathbf{h}}_s, \mathbf{a}_s), K_s^e(\cdot | \bar{\mathbf{h}}_s, \mathbf{a}_s))$, where $\text{TV}(\mu, \nu)$ denotes the total variation distance
 300 $\text{TV}(\mu, \nu) := \sup_{A \in \mathcal{A}} |\mu(A) - \nu(A)|$.

301 **Proposition 4.4** (Propagation-error bound and dominance). *Assume that the single-step AR tail
 302 operators are Lipschitz in total variation with factors $(1 + \lambda_j)$ (see Assumption E.7). Let P^* be the
 303 interventional law of $Y_{t+\tau}$ and $P^{\text{lat}}, P^{\text{AR}}$ the laws induced by latent and AR rollouts using $\{K_s^e\}$.
 304 Then, taking total variation over the marginal of $Y_{t+\tau}$,*

$$\text{TV}(P^*, P^{\text{lat}}) \leq \sum_{s=t}^{t+\tau-1} \varepsilon_s, \quad \text{TV}(P^*, P^{\text{AR}}) \leq \sum_{s=t}^{t+\tau-1} \varepsilon_s \prod_{j=s+1}^{t+\tau-1} (1 + \lambda_j). \tag{8}$$

309 Proof. App. E.10.

310 Our model inevitably makes small one-step errors in the conditional distributions. The key differ-
 311 ence is how these local errors are propagated. In the latent g -computation rollout, once the factual
 312 history is encoded, all future evolution happens in latent space and decoded predictions are never
 313 fed back; mathematically, the subsequent latent transitions are Markov and non-expansive in total
 314 variation, so each local error contributes at most additively to the final discrepancy. In a data-space
 315 autoregressive rollout, every decoded prediction is fed back through a powerful encoder to form
 316 the next context, and these encode-decode maps can enlarge discrepancies, so a small local error
 317 at a given time step can be amplified at later steps. Proposition 4.4 formalizes exactly this: both
 318 approaches share the same local approximation errors, but only the data-space rollout has this addi-
 319 tional error-amplification channel, which explains its worse long-horizon behavior.

321 4.3 NEURAL INSTANTIATION: THE G-LATENT MODEL

322 **Architecture.** We instantiate the abstract components $(f_\omega, \Gamma_\gamma, \kappa_\psi, p_\theta, p_0)$ with a history network, a
 323 latent GRU, and a conditional VAE. The *history network* f_ω is a multi-input transformer that maps

324 the observed history $\bar{\mathbf{h}}_t$ to an embedding $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$, following Melnychuk et al. (2022) (three
 325 streams for $\bar{\mathbf{x}}_t$, $\bar{\mathbf{a}}_{t-1}$, $\bar{\mathbf{y}}_t$ with cross-attention; details in App. B). The latent state update Γ_γ (Eq. 4)
 326 is implemented as a GRU, and the context map as $\mathbf{c}_{t,t'} = \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$, so that future
 327 contexts depend only on compact latent summaries rather than decoded observations. For the one-
 328 step conditionals over $\mathbf{l}_{t+t'} = (\mathbf{y}_{t+t'}, \mathbf{x}_{t+t'})$ we use a single conditional VAE, shared across t' , with
 329 VAE *encoder* and decoder

$$q_\phi(\mathbf{z}_{t,t'} \mid \mathbf{l}_{t+t'}, \mathbf{c}_{t,t'}), \quad p_\theta(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}),$$

330 and prior $p_0(\mathbf{z}_{t,t'}) = \mathcal{N}(\mathbf{0}, \mathbf{I})$. Outcomes $\mathbf{y}_{t+t'}$ are modeled with the ALD-mixture parameterization
 331 of An & Jeon (2023) (DistVAE), extended here to time series with sequential treatments, while
 332 covariates $\mathbf{x}_{t+t'}$ use Gaussian heads; see below.

333 **Training objective and implementation.** We share one conditional VAE across steps $t' \in$
 334 $\{1, \dots, \tau\}$ and optimize a joint per-step objective. Given the context $\mathbf{c}_{t,t'}$, the VAE encoder
 335 outputs $\mathbf{z}_{t,t'} \sim q_\phi(\mathbf{z}_{t,t'} \mid \mathbf{l}_{t+t'}, \mathbf{c}_{t,t'})$, and we update the latent state via Eq. 4. The decoder
 336 $p_\theta(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$ is parameterized by a *shared trunk* T_θ followed by two heads: an
 337 outcome head $D_\theta^{(y)}$ and a covariate head $D_\theta^{(x)}$. Let $\mathbf{w}_{t,t'} = T_\theta(\mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$ and

$$\hat{\mathbf{q}}_{\alpha,t,t'} = D_\theta^{(y)}(\mathbf{w}_{t,t'}, \alpha), \quad (\hat{\mu}_{t,t'}, \hat{\sigma}_{t,t'}^2) = D_\theta^{(x)}(\mathbf{w}_{t,t'}),$$

338 where $\alpha \in (0, 1)^{d_y}$ collects per-outcome quantile levels. We implement $D_\theta^{(y)}$ as d_y scalar branches
 339 and draw K vectors $\{\alpha^{(k)}\}_{k=1}^K$ with i.i.d. entries $\alpha_j^{(k)} \sim \text{Unif}(0, 1)$. The per-step reconstruction
 340 loss is

$$\mathcal{L}_{\text{rec}}(t, t') = \sum_{j=1}^{d_y} \frac{1}{K} \sum_{k=1}^K \rho_{\alpha_j^{(k)}}(y_{t+t',j} - \hat{q}_{\alpha_j^{(k)}, t, t', j}) + \frac{1}{2} \left\| \frac{\mathbf{x}_{t+t'} - \hat{\mu}_{t,t'}}{\hat{\sigma}_{t,t'}} \right\|_2^2 + \frac{1}{2} \mathbf{1}^\top \log \hat{\sigma}_{t,t'}^2, \quad (9)$$

341 where $\rho_\alpha(u) = (\alpha - \mathbf{1}\{u < 0\})u$ is the pinball loss and $(\hat{\mu}_{t,t'}, \hat{\sigma}_{t,t'}^2)$ are the Gaussian parameters
 342 for $\mathbf{x}_{t+t'}$. The KL term is $\mathcal{L}_{\text{KL}}(t, t') = \text{KL}(q_\phi(\mathbf{z}_{t,t'} \mid \cdot) \parallel \mathcal{N}(\mathbf{0}, \mathbf{I}))$. This corresponds to a conditional
 343 VAE with an ALD-mixture outcome decoder (An & Jeon, 2023); integrating over α recovers a
 344 CRPS-type reconstruction term, which encourages well-calibrated, flexible predictive distributions
 345 beyond Gaussian heads (see App. C for details). In our setting, using the ALD mixture for \mathbf{y} im-
 346 proves distributional performance but increases decoder complexity, so we use it only for outcomes
 347 and keep a simpler Gaussian head for covariates \mathbf{x} , where the additional expressivity does not off-
 348 set the extra compute. Predictive uncertainty arises both from the sampled latent path (capturing
 349 temporal and cross-outcome dependence) and from the outcome head, which plays the role of the
 350 likelihood noise model, analogous to decoder noise in a Gaussian VAE.

351 The history network f_ω is high-capacity, and the VAE objective alone can be minimized even if
 352 \mathbf{r}_t carries little predictive signal (the decoder may partly ignore it). To avoid such degenerate con-
 353 figurations, we add an auxiliary one-step prediction head $\hat{\mathbf{y}}_{t+1} = U_\eta(\mathbf{r}_t, \mathbf{a}_t)$ with MSE loss \mathcal{L}_{aux}
 354 (Eq. 10), used purely as a regularizer to make \mathbf{r}_t predictive of \mathbf{y}_{t+1} . The total loss over a mini-batch
 355 \mathcal{B} is

$$\mathcal{L} = \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} \sum_{t=1}^{T^{(i)}-1} \left[\sum_{t'=1}^{\tau} m_{t,t'}^{(i)} (\mathcal{L}_{\text{rec}}^{(i)}(t, t') + \beta \mathcal{L}_{\text{KL}}^{(i)}(t, t')) + \lambda_{\text{aux}} m_{t,1}^{(i)} \mathcal{L}_{\text{aux}}^{(i)}(t) \right], \quad (10)$$

356 with masks $m_{t,t'}^{(i)} = \mathbf{1}\{t + t' \leq T^{(i)}\}$. In practice, we found it helpful to *warm start* the history
 357 network by first optimizing only \mathcal{L}_{aux} for a small number of epochs, and then training the full
 358 objective in Eq. 10. This implementation choice affects how the parameters are learned but does
 359 not change the latent g-computation estimator of Section 4.1. We also reweight the two terms in
 360 Eq. 9 to give more importance to outcome modeling. Hyperparameters are selected via lightweight
 361 tuning on factual-validation sets, guided by distributional metrics and KL-capacity diagnostics; for
 362 the transformer we adopt the architecture and base hyperparameters of Melnychuk et al. (2022).

363 **Inference and sampling cost.** At test time, we apply Algorithm 1 with the learned parameters. For
 364 a given anchor time t and treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$, we compute the history embedding $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$
 365 once, then roll out the latent state and decoder as in Section 4.1. Because decoded observations
 366 are never fed back, the inner loop consists only of GRU updates and decodes and vectorizes over

378 M Monte Carlo paths with a shared \mathbf{r}_t . As discussed in Section 4.1 and Corollary 4.3, we can
 379 decode at all steps (`scope=all`) or only at a subset $S \subseteq \{1, \dots, \tau\}$ (e.g., `scope=last` for $S =$
 380 $\{\tau\}$) without changing the underlying interventional law, so we pay decoder cost only at horizons
 381 of interest. For M MC samples and horizon τ , the cost is $\mathcal{O}(\text{cost}(f_\omega) + M[\tau(\text{cost}(\text{GRU}^{(z)}) +$
 382 $\text{cost}(\kappa_\psi)) + |S| \text{cost}(D_\theta^{(y)})])$, where $|S| \leq \tau$ is the number of decoded steps and $\text{cost}(f_\omega)$ is paid
 383 once. By contrast, a data-space rollout has cost $\mathcal{O}(\text{cost}(f_\omega) + M\tau[\text{cost}(\text{GRU}^{(L)}) + \text{cost}(\kappa_\psi) +$
 384 $\text{cost}(D_\theta^{(x,y)})])$, since all steps and both X and Y must be decoded, and a full autoregressive model
 385 with decoder scales as $\mathcal{O}(M\tau[\text{cost}(f_\omega) + \text{cost}(D_\theta^{(x,y)})])$ (G-Net is of this type, but uses a hold-
 386 out error set instead of a decoder). Overall, our model reduces sample cost by (i) computing f_ω
 387 once and reusing it across M and all τ steps, enabling a high-capacity transformer only for the
 388 up-to- t sequence; (ii) decoding selectively so the D_θ term scales with $|S|$ (e.g., $|S|=1$ for `last`);
 389 (iii) decoding only $D_\theta^{(y)}$ and skipping $D_\theta^{(x)}$ at inference; and (iv) updating the GRU in latent space
 390 ($\text{GRU}^{(z)}$) instead of data space ($\text{GRU}^{(L)}$), which can yield gains when $d_z \ll d_L$.
 391
 392
 393
 394

5 EVALUATION

397 **Datasets.** Following common practice in benchmarking for POs inference (Bica et al.; Melnychuk
 398 et al., 2022), we make use of a semi-synthetic dataset for validating our approach, as it allows
 399 to compute ground truth POs. Additionally, we also use a real-world dataset to demonstrate the
 400 practical applicability of our approach. These datasets were selected because they have a consider-
 401 able number of covariates to adjust for, which is the type of setting for which our model can be
 402 more useful. *Semi-synthetic*: from ICU data (Johnson et al., 2016), we generate high-dimensional,
 403 long-range trajectories with treatment effects and endogenous/exogenous dependencies following
 404 Melnychuk et al. (2022); Schulam & Saria (2017); confounding is controllable and ground-truth
 405 POs are known. We detected violations of the positivity assumption in the original form of this
 406 dataset, presented in Melnychuk et al. (2022). Despite having become a standard benchmark, the
 407 aforementioned positivity violations make it unsuitable for evaluation of methods with the stand-
 408 ard causal assumptions. For this reason, we make several modifications to avoid this problem.
 409 We detail the detected problems in the original form of the dataset and the changes we make in F.
 410 *Real-world*: a fully observational benchmark from MIMIC-III using the same cohort definition and
 411 preprocessing as the semi-synthetic setup (sampling grid, variable definitions, imputation, and dis-
 412 crete action categories per Melnychuk et al., 2022); lacking ground-truth counterfactuals, evaluation
 413 targets predictive quality of observational next steps. Variables include standard ICU vitals/labs and
 414 intervention-derived action indicators. We refer to App. F for more details about both datasets.

415 **Baselines.** To evaluate our model, we use several baselines that handle aleatoric uncertainty and
 416 deliver distributional estimates. We use G-Net (Li et al., 2021) as an alternative implementation of
 417 the g-formula and, for better comparability, its extension Transformer G-Net (Xiong et al., 2024),
 418 which we implement with the same multi-input transformer architecture used in G-Latent. To the
 419 best of our knowledge, these are the only previous works that estimate aleatoric uncertainty of indi-
 420 vidualized POs in a discrete setting. We also compare with Causal Transformer (CT) (Melnychuk
 421 et al., 2022): in its original form for point estimate metrics, and with two distributional adaptations:
 422 CT-Gaussian, with a Gaussian head, and CT-CRPS, with a CRPS head, analogous to G-Latent de-
 423 coder. Among the non-distributional models for individualized POs, we chose to adapt CT as it is
 424 a strong baseline and G-Latent shares its transformer-based processing of history data. As for our
 425 model, we present three variants apart from the one described in 4.3: G-Latent with a full Gaussian
 426 reconstruction, and two variants that perform the rollout in the data space: one with CRPS decoder
 427 and another one with full Gaussian decoder. We call these variants G-VAE, and D.S. accounts for
 428 data space. We specify the details in App. G. In continuous settings, we are aware of two works
 429 that estimate data distributions: Hess et al. and Mu et al. (2025). We exclude the former because it
 430 introduces a heavy machinery for epistemic uncertainty and continuous time processing that makes
 431 it very expensive to train, while its way to handle aleatoric uncertainty is a Gaussian head, which
 432 is already covered by CT-Gaussian. As for the latter, we exclude it because it addresses a slightly
 433 different setting (expert models) and because it was released over one month before the submission
 434 of this work, without available code.

432 Table 1: Results at selected steps $t' \in \{3, 5, 8, 11\}$ for the (new) semi-synthetic dataset. Metrics:
433 Energy Score (ES \downarrow) (per step and across steps), KDE-Loglikelihood (KDE-LL \uparrow), RMSE \downarrow , Cali-
434 bration MAE \downarrow .

Model	$t' = 3$			$t' = 5$			$t' = 8$			$t' = 11$			Global	
	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	Cal. MAE \downarrow
G-Net	0.39 \pm 0.04	-1.27 \pm 0.17	0.64 \pm 0.07	0.51 \pm 0.05	-1.74 \pm 0.21	0.81 \pm 0.09	0.63 \pm 0.07	-2.18 \pm 0.25	0.98 \pm 0.11	0.64 \pm 0.08	-2.45 \pm 0.28	1.09 \pm 0.12	1.85 \pm 0.20	6.29 \pm 1.35
Transformer G-Net	0.40 \pm 0.05	-1.35 \pm 0.21	0.66 \pm 0.08	0.50 \pm 0.07	-1.69 \pm 0.31	0.80 \pm 0.13	0.58 \pm 0.11	-2.01 \pm 0.45	0.92 \pm 0.10	0.64 \pm 0.14	-2.24 \pm 0.56	1.00 \pm 0.23	1.71 \pm 0.11	6.97 \pm 1.26
CT (CRPS)	0.32 \pm 0.07	-1.00 \pm 0.30	0.58 \pm 0.11	0.41 \pm 0.07	-1.40 \pm 0.34	0.71 \pm 0.10	0.50 \pm 0.07	-1.87 \pm 0.35	0.84 \pm 0.10	0.57 \pm 0.07	-2.22 \pm 0.36	0.92 \pm 0.10	1.52 \pm 0.23	13.14 \pm 2.55
CT (Gaussian)	0.30 \pm 0.07	-0.91 \pm 0.31	0.54 \pm 0.13	0.37 \pm 0.08	-1.17 \pm 0.35	0.64 \pm 0.14	0.44 \pm 0.09	-1.44 \pm 0.38	0.74 \pm 0.14	0.49 \pm 0.09	-1.64 \pm 0.38	0.81 \pm 0.14	1.35 \pm 0.29	7.88 \pm 1.76
CT	0.43 \pm 0.10	0.53 \pm 0.12	0.60 \pm 0.13	...	0.65 \pm 0.13
D.S. G-VAE (Gaussian)	0.49 \pm 0.04	-2.26 \pm 0.12	0.54 \pm 0.09	0.58 \pm 0.06	-2.56 \pm 0.16	0.66 \pm 0.11	0.64 \pm 0.07	-2.72 \pm 0.18	0.76 \pm 0.13	0.67 \pm 0.07	-2.78 \pm 0.18	0.83 \pm 0.13	2.01 \pm 0.20	14.99 \pm 0.86
D.S. G-VAE (CRPS)	0.28 \pm 0.05	-0.89 \pm 0.25	0.49 \pm 0.10	0.35 \pm 0.06	-1.14 \pm 0.29	0.59 \pm 0.12	0.42 \pm 0.07	-1.40 \pm 0.29	0.69 \pm 0.12	0.47 \pm 0.06	-1.58 \pm 0.26	0.76 \pm 0.12	1.28 \pm 0.21	5.48 \pm 3.06
G-Latent (Gaussian)	0.38 \pm 0.04	-1.70 \pm 0.14	0.53 \pm 0.09	0.42 \pm 0.05	-1.80 \pm 0.16	0.61 \pm 0.11	0.46 \pm 0.06	-1.90 \pm 0.18	0.69 \pm 0.12	0.48 \pm 0.06	-1.95 \pm 0.18	0.73 \pm 0.12	1.51 \pm 0.18	10.14 \pm 1.36
G-Latent (CRPS)	0.29 \pm 0.05	-0.95 \pm 0.21	0.51 \pm 0.10	0.35 \pm 0.06	-1.18 \pm 0.26	0.60 \pm 0.12	0.40 \pm 0.07	-1.37 \pm 0.29	0.68 \pm 0.13	0.43 \pm 0.08	-1.50 \pm 0.29	0.73 \pm 0.13	1.25 \pm 0.23	2.95 \pm 1.37

442 Table 2: Results at selected steps $t' \in \{2, 3, 5, 6\}$ for the real-world dataset. Metrics: Energy Score
443 (ES \downarrow) (per step and across steps), KDE-Loglikelihood (KDE-LL \uparrow), and RMSE \downarrow .

Model	$t' = 2$			$t' = 3$			$t' = 5$			$t' = 6$			Global	
	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	
G-Net	5.32 \pm 0.08	-3.92 \pm 0.05	11.84 \pm 0.24	5.82 \pm 0.08	-4.11 \pm 0.05	12.83 \pm 0.29	6.98 \pm 0.09	-4.55 \pm 0.07	14.05 \pm 0.30	7.44 \pm 0.11	-4.83 \pm 0.04	14.23 \pm 0.29	18.35 \pm 0.33	
Transformer G-Net	5.28 \pm 0.06	-3.89 \pm 0.06	10.90 \pm 0.30	5.84 \pm 0.08	-4.06 \pm 0.08	11.67 \pm 0.26	6.47 \pm 0.08	-4.30 \pm 0.06	12.96 \pm 0.32	6.90 \pm 0.08	-4.48 \pm 0.04	13.21 \pm 0.29	16.70 \pm 0.23	
CT (CRPS)	4.92 \pm 0.06	-3.81 \pm 0.06	10.10 \pm 0.20	5.39 \pm 0.08	-3.94 \pm 0.08	10.53 \pm 0.26	5.77 \pm 0.08	-4.08 \pm 0.04	10.75 \pm 0.29	5.86 \pm 0.07	-4.19 \pm 0.06	10.91 \pm 0.28	14.61 \pm 0.27	
CT (Gaussian)	5.25 \pm 0.06	-3.92 \pm 0.06	10.41 \pm 0.29	5.71 \pm 0.08	-4.04 \pm 0.07	10.74 \pm 0.29	6.15 \pm 0.07	-4.18 \pm 0.06	11.01 \pm 0.34	6.34 \pm 0.08	-4.24 \pm 0.07	11.25 \pm 0.30	15.55 \pm 0.23	
CT	...	0.90 \pm 0.23	...	0.57 \pm 0.24	...	10.16 \pm 0.27	10.35 \pm 0.31	
D.S. G-VAE (Gaussian)	5.51 \pm 0.08	-3.90 \pm 0.06	9.58 \pm 0.25	5.99 \pm 0.08	-3.98 \pm 0.06	10.29 \pm 0.22	6.34 \pm 0.06	-4.03 \pm 0.05	10.88 \pm 0.26	6.44 \pm 0.07	-4.04 \pm 0.05	11.04 \pm 0.29	15.98 \pm 0.23	
D.S. G-VAE (CRPS)	4.89 \pm 0.08	-3.82 \pm 0.06	9.40 \pm 0.22	5.36 \pm 0.08	-3.92 \pm 0.05	10.09 \pm 0.25	5.70 \pm 0.07	-3.99 \pm 0.06	10.63 \pm 0.29	5.82 \pm 0.06	-4.04 \pm 0.06	10.79 \pm 0.30	14.38 \pm 0.19	
G-Latent (Gaussian)	5.27 \pm 0.06	-3.85 \pm 0.06	9.42 \pm 0.23	5.64 \pm 0.08	-3.89 \pm 0.05	10.09 \pm 0.23	5.96 \pm 0.07	-3.94 \pm 0.04	10.64 \pm 0.19	6.07 \pm 0.07	-3.95 \pm 0.06	10.80 \pm 0.25	15.21 \pm 0.26	
G-Latent (CRPS)	4.85 \pm 0.05	-3.79 \pm 0.06	9.23 \pm 0.20	5.25 \pm 0.08	-3.88 \pm 0.05	9.79 \pm 0.24	5.60 \pm 0.09	-3.94 \pm 0.05	10.36 \pm 0.29	5.72 \pm 0.06	-3.96 \pm 0.06	10.55 \pm 0.28	14.23 \pm 0.23	

452 **Metrics.** Our model produces MC samples at each prediction step. We evaluate with: *RMSE of the*
453 *predictive mean*, computed from the average of MC samples at each step (lower is better); *Energy*
454 *Score (ES)*, a strictly proper multivariate scoring rule that reduces to CRPS in the univariate case
455 and assesses distributional fit. We report it per step and over the full trajectory to capture temporal
456 coherence (lower is better); and *KDE log-likelihood (KDE-LL)*, the log-likelihood of the observed
457 outcome under a Gaussian kernel density estimate fit to the model’s samples, reflecting density fit
458 (higher is better). After trying over ten bandwidths for each dataset and baseline, we selected the one
459 with general better results to report here. For the semi-synthetic dataset, we report results for two ad-
460 dditional bandwidth (see App. J). In general, the best bandwidths provided better results consistently
461 across models. For the semi-synthetic dataset, We also assess *calibration* via *quantile coverage*: for
462 $q \in \{0.1, \dots, 0.9\}$ we compute, per step and per outcome dimension (and aggregated across steps),
463 the fraction of test outcomes below the MC-estimated q -quantile (ideal coverage equals q). As a
464 scalar summary we report *Calibration MAE*, the mean absolute gap between empirical and nominal
465 coverage averaged over quantiles, dimensions, and steps (lower is better). To obtain the metrics,
466 we used 50 and 40 MC samples for the semi-synthetic and the real-world dataset, respectively. See
467 App. H for more details on the metrics.

468 **Results.** We ran all experiments in AWS SageMaker on an ml.g5.4xlarge instance (A10G GPU,
469 24 GiB VRAM). We report selected steps in Table 1 (semi-synthetic, modified) and Table 2 (real-
470 world), with full results—and the original semi-synthetic benchmark—in App. J. Semi-synthetic
471 runs use five random seeds; real-world runs use four; intervals denote standard deviations.¹ Across
472 both datasets, **G-Latent** attains the strongest *distributional* performance, especially at larger hori-
473 zons. On semi-synthetic data, *G-VAE-CRPS* remains competitive with *G-Latent-CRPS*—showing
474 small ES gaps overall and occasional wins at short horizons—whereas among the Gaussian variants
475 the gap between *G-Latent* and *G-VAE* is pronounced: Gaussian heads are more error-prone, and
476 the latent rollout reduces accumulation error. KDE log-likelihood consistently favors **G-Latent** at
477 large steps (across all tested bandwidths). On the real-world cohort, *G-Latent-CRPS* is best at ev-
478 ery reported step and globally. For calibration on the semi-synthetic benchmark, *G-Latent-CRPS*
479 achieves the lowest Calibration MAE by a clear margin, while Gaussian variants fare markedly
480 worse. In App. J we show extensive quantile coverage tables. Regarding other baselines, CT with
481 Gaussian/CRPS heads trails the latent models on distributional metrics, while the point-estimate CT
482 attains the lowest RMSE (as expected for a point forecaster); G-Net and Transformer G-Net lag fur-
483 ther behind on ES and KDE-LL. Overall, *G-Latent-CRPS* provides the best distributional metrics
484 at long horizons while remaining competitive on point accuracy, and it clearly outperforms prior
485 g-computation-based models.

¹See App. J for complete tables and diagnostics.

486 We measure end-to-end test-set inference time on the semi-
 487 synthetic dataset (50 MC samples; 11 projection-horizon
 488 steps). Table 3 reports the results: decoding all steps with
 489 G-Latent–CRPS takes 00:19:27 (1,167 s \pm 12 s), while de-
 490 coding only the last step takes 00:07:11 (431 s \pm 5 s)—an
 491 \approx 63% reduction that is valuable when only a few hori-
 492 zons are needed, since non-latent rollouts must decode every
 493 step. For G-VAE–CRPS, inference time is 00:25:42 (1,542 s
 494 \pm 12 s), about 32% slower than G-Latent–CRPS (all steps).
 495 This gap stems from our decoupled decoder, which allows G-
 496 Latent–CRPS to decode outcomes without covariates. In our
 497 implementation, the outcome and covariate decoders share
 498 three layers (App. D); further decoupling could yield addi-
 499 tional gains. The Gaussian head yields similar wall-clock for G-Latent—00:20:16 (1,216 s \pm 15 s)
 500 for all steps and 00:07:26 (446 s \pm 8 s) for last-step decoding—and 00:20:05 (1,205 s \pm 11 s) for
 501 G-VAE (there is no covariate decoupling in the Gaussian head models). Among other baselines,
 502 Transformer G-Net and CT-CRPS/CT-Gaussian are substantially slower at 01:03:21 (3,801 s \pm
 503 36 s), 00:59:25 (3,565 s \pm 29 s), and 00:53:08 (3,188 s \pm 19 s), respectively, while G-Net is faster
 504 at 00:05:45 (345 s \pm 5 s). For all the baselines, we fully tensorize and cache recurrent state (e.g.,
 505 Transformer hidden states in Transformer G-Net and CT-CRPS/Gaussian), so each step only pro-
 506 cesses the last MC prediction rather than recomputing the entire history. In summary, all full trans-
 507 former-based models exceed 50 minutes per test set, whereas G-Latent (and its variants) substan-
 508 tially reduces inference time by using the transformer only to encode the history up-to- t' , then up-
 509 dating the representation during the MC rollout with a lightweight GRU. Our tensorized and cached
 510 implementation of G-Net achieves very low inference times because it uses a lightweight RNN to
 511 process data and, unlike G-Latent, has no decoder—it injects residual noise. However, this reduces
 512 its expressivity and adaptability to particular data distributions.

6 CONCLUSIONS AND LIMITATIONS

515 In this work, we introduce G-Latent, a novel method for distributional estimation of individual-
 516 ized POs under time-varying treatment effects for discrete settings, with identifiability guarantees
 517 through g-computation in the latent space. We demonstrate the general efficacy of our approach,
 518 both theoretically and experimentally. Also, we show that our method is efficient at sampling com-
 519 pared with other variants that perform g-computation in the data-space. We identify two potential
 520 limitations: the first is related to the latent factorization in eq. 6, fundamental for G-Latent. This
 521 assumption would be violated, for example, under posterior collapse (Lucas et al., 2019), which is
 522 relatively common in VAE training and prevents latent representations from properly representing
 523 data. We did not observe this problem in the experiments, but it is important to be careful with
 524 that. On the other hand, another potential limitation comes from the CRPS decoder; as An & Jeon
 525 (2023) discuss, the ALD-decoder assumes that the different elements of $\mathbf{Y}_{t+t'}$ (if multivariate) are
 526 independent given \mathbf{z} . If the assumption fails, cross-dimensional dependence may remain unmod-
 527 eled. However, neither DistVAE nor us empirically observe this problem (G-Latent has strong ES
 528 metrics). Finally, our focus in this work is aleatoric uncertainty; epistemic uncertainty is orthogonal
 529 and can be added with MC dropout or deep ensembles, or more formally via Bayesian priors.

530 We restrict attention to g-computation-based estimators rather than IPTW/MSM-style generative
 531 baselines (e.g., Wu et al., 2024). In principle, IPTW could be adapted to our conditional, trajectory-
 532 level estimands, but would require high-dimensional propensity models (or conditional treatment
 533 densities for continuous treatments) and weighted conditional density estimation, which can lead to
 534 unstable importance weights in long-horizon, high-dimensional settings. Designing and evaluating
 535 IPTW/MSM-style generative models for individualized distributional potential outcomes remains
 an interesting direction for future work.

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Table 3: Test-set inference time on the semi-synthetic dataset (50 MC samples; 11 projection-horizon steps) (hh:mm:ss).

Method	hh:mm:ss
G-Latent (CRPS) [all]	00:19:27 \pm 12s
G-Latent (CRPS) [last]	00:07:11 \pm 05s
G-Latent (Gaussian) [all]	00:20:16 \pm 15s
G-Latent (Gaussian) [last]	00:07:26 \pm 08s
G-VAE (CRPS)	00:25:42 \pm 12s
G-VAE (Gaussian)	00:20:05 \pm 11s
Transformer G-Net	01:03:21 \pm 36s
G-Net	00:05:45 \pm 05s
CT-CRPS	00:59:25 \pm 29s
CT-Gaussian	00:53:08 \pm 19s

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674

A ASSUMPTIONS FOR CAUSAL IDENTIFICATION

675 We work within the potential outcomes paradigm (Rubin, 2005) and its extension to temporal treat-
 676 ments and outcomes (Robins et al., 2000), a setup also adopted by prior sequence models for treat-
 677 ment effect inference (e.g., Lim, 2018; Bica et al.). In this framework, identification of counterfac-
 678 tional distributions over time (and, in particular, the τ -step conditional mean from Eq. (1)) relies on
 679 three standard conditions on the data-generating process.

680 **Assumption A.1 (Consistency).** For any fixed treatment history $\bar{\mathbf{a}}_t$, if the realized actions satisfy
 681 $\bar{\mathbf{A}}_t = \bar{\mathbf{a}}_t$, then

$$\mathbf{Y}_{t+1}[\bar{\mathbf{a}}_t] = \mathbf{Y}_{t+1}.$$

682 That is, under the actually received treatment sequence, the relevant potential outcome coincides
 683 with the observed one.

684 **Assumption A.2 (Sequential Overlap/Positivity).** For any history value $\bar{\mathbf{h}}_t$ in the support of $\bar{\mathbf{H}}_t$,
 685 each admissible action has positive probability:

$$0 < p(\mathbf{A}_t = \mathbf{a}_t \mid \bar{\mathbf{H}}_t = \bar{\mathbf{h}}_t) < 1 \quad \text{whenever} \quad p(\bar{\mathbf{H}}_t = \bar{\mathbf{h}}_t) > 0.$$

686 **Assumption A.3 (Sequential Ignorability / No Unmeasured Confounding).** Conditioning on
 687 the observed history renders the current action as-if randomized with respect to the next-step poten-
 688 tial outcome:

$$\forall t \text{ and } \forall \bar{\mathbf{a}}_{t:t+\tau-1} : \mathbf{A}_t \perp\!\!\!\perp (\bar{\mathbf{L}}_{t+1:t+\tau}[\bar{\mathbf{a}}_{t:t+\tau-1}], \bar{\mathbf{Y}}_{t+1:t+\tau}[\bar{\mathbf{a}}_{t:t+\tau-1}]) \mid \bar{\mathbf{H}}_t.$$

689 **Corollary A.4 (g-computation; Robins & Hernan, 2008).** Under A.1–A.3, the τ -step-ahead con-
 690 ditional mean under a fixed intervention path $\bar{\mathbf{a}}_{t:t+\tau-1}$ is identified by the longitudinal g -formula.

702 B MULTI-INPUT TRANSFORMER

704 **Scope.** This appendix details the *encoder* we use to compute the history embedding $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$
 705 from the factual history $\bar{\mathbf{h}}_t = \{\bar{\mathbf{X}}_t, \bar{\mathbf{A}}_{t-1}, \bar{\mathbf{Y}}_t, \mathbf{V}\}$. It follows the multi-input transformer design
 706 of Melnychuk et al. (2022) (three streams with cross-attention and shared relative positional en-
 707 codings), but we do *not* use their balancing loss and we *never* feed model predictions back into
 708 the transformer. The output of this encoder is a single fused representation \mathbf{r}_t that our model uses
 709 downstream (Sec. 4.3).

710 B.1 INPUTS AND TOKENIZATION

712 Let $b = 1, \dots, B$ index transformer blocks and d_h the model width. For the first block, we map
 713 each sequence to hidden states via time-shared linear layers:

715 $\mathbf{A}_{1:t}^0 = \text{Linear}_A(\bar{\mathbf{A}}_{t-1}), \quad \mathbf{X}_{1:t}^0 = \text{Linear}_X(\bar{\mathbf{X}}_t), \quad \mathbf{Y}_{1:t}^0 = \text{Linear}_Y(\bar{\mathbf{Y}}_t), \quad \tilde{\mathbf{V}} = \text{Linear}_V(\mathbf{V}),$
 716 where $\bar{\mathbf{A}}_{t-1} = (\mathbf{A}_1, \dots, \mathbf{A}_{t-1}, \mathbf{0})$ is a left-shifted treatment stream aligned with our start-of-
 717 interval indexing (decision \mathbf{A}_j precedes $(\mathbf{Y}_{j+1}, \mathbf{X}_{j+1})$). Subsequent blocks receive the previous
 718 block's outputs.

719 We denote the stream-specific hidden sequences at block b by $\mathbf{A}_{1:t}^b$, $\mathbf{X}_{1:t}^b$, and $\mathbf{Y}_{1:t}^b$ ($\in \mathbb{R}^{t \times d_h}$).

721 B.2 MASKED SELF-ATTENTION WITH RELATIVE POSITIONAL ENCODINGS

723 Each stream applies masked multi-head self-attention (causal mask so a position i only attends to
 724 $j \leq i$) with *relative* positional encodings (RPE). For head dimension d_{qk} , attention at position i is

$$726 \text{Attn}_i(Q, K, V) = \sum_{j=1}^t \alpha_{ij} (V_j + a_{ij}^V), \quad \alpha_{ij} = \text{softmax}_j \left(\frac{Q_i^\top (K_j + a_{ij}^K)}{\sqrt{d_{qk}}} \right), \quad (11)$$

$$728 a_{ij}^V = w_{\text{clip}(j-i, \ell_{\max})}^V, \quad a_{ij}^K = w_{\text{clip}(j-i, \ell_{\max})}^K, \quad \text{clip}(x, \ell_{\max}) = \max\{-\ell_{\max}, \min\{\ell_{\max}, x\}\},$$

730 with trainable $w_\ell^V, w_\ell^K \in \mathbb{R}^{d_{qk}}$ for $\ell \in \{-\ell_{\max}, \dots, 0\}$. These Toeplitz-structured encodings
 731 depend only on relative distance and are shared across blocks and streams. Layer normalization
 732 and residual connections wrap the attention sublayer, and a position-wise feed-forward network
 733 $\text{FF}(h) = \text{Linear}(\text{ReLU}(\text{Linear}(h)))$ follows, again with residual+LN.

734 B.3 CROSS-ATTENTION BETWEEN STREAMS AND STATIC COVARIATES

736 To couple signals across modalities, each block augments self-attention with *cross-attentions* be-
 737 tween the three streams. Using tildes for post-self-attention states and writing $\text{MHA}(Q, K, V)$ for
 738 multi-head attention,

$$740 \tilde{\mathbf{A}}_X^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{A}}^{b-1}), K(\mathbf{X}^{b-1}), V(\mathbf{X}^{b-1})) + \tilde{\mathbf{A}}^{b-1} \right), \quad (12)$$

$$742 \tilde{\mathbf{A}}_Y^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{A}}^{b-1}), K(\mathbf{Y}^{b-1}), V(\mathbf{Y}^{b-1})) + \tilde{\mathbf{A}}^{b-1} \right), \quad (13)$$

$$744 \tilde{\mathbf{X}}_A^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{X}}^{b-1}), K(\mathbf{A}^{b-1}), V(\mathbf{A}^{b-1})) + \tilde{\mathbf{X}}^{b-1} \right), \quad (14)$$

$$746 \tilde{\mathbf{X}}_Y^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{X}}^{b-1}), K(\mathbf{Y}^{b-1}), V(\mathbf{Y}^{b-1})) + \tilde{\mathbf{X}}^{b-1} \right), \quad (15)$$

$$748 \tilde{\mathbf{Y}}_X^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{Y}}^{b-1}), K(\mathbf{X}^{b-1}), V(\mathbf{X}^{b-1})) + \tilde{\mathbf{Y}}^{b-1} \right), \quad (16)$$

$$750 \tilde{\mathbf{Y}}_A^{b-1} = \text{LN} \left(\text{MHA}(Q(\tilde{\mathbf{Y}}^{b-1}), K(\mathbf{A}^{b-1}), V(\mathbf{A}^{b-1})) + \tilde{\mathbf{Y}}^{b-1} \right). \quad (17)$$

751 We then pool the two cross-attended views per stream and inject static covariates at every time step:

$$752 \tilde{\mathbf{A}}^{b-1} = \tilde{\mathbf{A}}_X^{b-1} + \tilde{\mathbf{A}}_Y^{b-1} + \mathbf{1}\tilde{\mathbf{V}}^\top, \quad \tilde{\mathbf{X}}^{b-1} = \tilde{\mathbf{X}}_A^{b-1} + \tilde{\mathbf{X}}_Y^{b-1} + \mathbf{1}\tilde{\mathbf{V}}^\top, \quad (18)$$

$$753 \tilde{\mathbf{Y}}^{b-1} = \tilde{\mathbf{Y}}_X^{b-1} + \tilde{\mathbf{Y}}_A^{b-1} + \mathbf{1}\tilde{\mathbf{V}}^\top, \quad (19)$$

755 followed by parallel FF+residual+LN sublayers to yield $\mathbf{A}^b, \mathbf{X}^b, \mathbf{Y}^b$. Treatments remain left-shifted
 throughout (so treatment token at index i aligns with covariate/outcome tokens at $i+1$).

756 B.4 FUSION TO A SINGLE HISTORY EMBEDDING \mathbf{r}_t
757758 After the final block B , we fuse the three streams by element-wise averaging at each time $i \leq t$,
759 then project with a linear layer and ELU:

760
$$\tilde{\Phi}_i = \frac{1}{3}(\mathbf{A}_{i-1}^B + \mathbf{X}_i^B + \mathbf{Y}_i^B), \quad \Phi_i = \text{ELU}(\text{Linear}(\tilde{\Phi}_i)), \quad \mathbf{r}_t := \Phi_t \in \mathbb{R}^{d_r}.$$

761

762 We use only the factual $\{\mathbf{X}_{1:t}, \mathbf{A}_{1:t-1}, \mathbf{Y}_{1:t}\}$ to build \mathbf{r}_t ; predicted outcomes are *never* fed back into
763 the encoder.764 **Remarks.** (i) All attention modules use the causal mask and the same RPE as in Eq. 11. (ii) Static
765 covariates \mathbf{V} are injected at every block/time step via $\tilde{\mathbf{V}}$. (iii) Dropout is applied after linear layers
766 in attention and feed-forward sublayers.769 C DISTVAE-STYLE LOSS: DERIVATION AND DISCUSSION
770771 We adapt the continuous-variable objective of An & Jeon (2023) to our setting (ignoring categorical
772 variables). Let $x = (x_1, \dots, x_p)$ denote continuous observations (here, $x \equiv \mathbf{y}$) and z the latent.
773 DistVAE assumes an *ALD* (asymmetric Laplace) decoder *mixed* over a quantile level $\alpha \in (0, 1)$:

774
$$p(x; \theta, \beta) = \iint p(x | z, \alpha; \theta, \beta) p(z) p(\alpha) d\alpha dz, \quad p(x | z, \alpha; \theta, \beta) = \prod_{j=1}^p p(x_j | z, \alpha; \theta_j, \beta), \quad (20)$$

775

776 where, for each coordinate,

777
$$p(x_j | z, \alpha; \theta_j, \beta) = \frac{\alpha(1-\alpha)}{\beta} \exp\left(-\rho_\alpha\left(\frac{x_j - D_j(\alpha, z; \theta_j)}{\beta}\right)\right), \quad \rho_\alpha(u) = (\alpha - \mathbb{I}\{u < 0\}) u. \quad (21)$$

778

779 Here $D_j(\alpha, z; \theta_j)$ is the conditional *quantile function* (ALD location)², $\beta > 0$ is a scale constant,
780 and ρ_α is the pinball loss.781 **Assumption 1 (DistVAE).** (i) $\{x_j\}$ are conditionally independent given z ; (ii) (discrete variables
782 independent of α ; not used here); (iii) $\alpha \perp z$. Item (i) is the usual VAE factorization; (iii) treats α
783 as a prior (no $q(\alpha | x)$), which is key to the proper-scoring-rule objective below.784 C.1 FINITE- K NEGATIVE ELBO (COMPOSITE QUANTILE)785 Approximate the α -integral by a uniform grid $\alpha_k = \frac{k}{K}$, $k = 1, \dots, K$, with $p(\alpha_k) = \frac{1}{K}$, and
786 introduce $q_\phi(z | x)$. A Jensen step yields, up to additive constants independent of (θ, ϕ) ,

787
$$-\text{ELBO}_K(\theta, \phi) = \mathbb{E}_{q_\phi(z | x)} \left[\frac{1}{K} \sum_{k=1}^K \sum_{j=1}^p \rho_{\alpha_k}(x_j - D_j(\alpha_k, z; \theta_j)) \right] + \beta \text{KL}(q_\phi(z | x) \| p(z)) + C_K, \quad (22)$$

788

789 so the reconstruction is a *composite quantile* (average ALD NLL across $\{\alpha_k\}$).790 C.2 LIMIT $K \rightarrow \infty$: CRPS OBJECTIVE AND DISTVAE LOSS791 Under mild integrability/continuity in α ,

792
$$\lim_{K \rightarrow \infty} \frac{1}{K} \sum_{k=1}^K \rho_{\alpha_k}(x_j - D_j(\alpha_k, z; \theta_j)) = \int_0^1 \rho_\alpha(x_j - D_j(\alpha, z; \theta_j)) d\alpha, \quad (23)$$

793

794
$$\lim_{K \rightarrow \infty} \frac{1}{K} \sum_{k=1}^K \log \alpha_k(1 - \alpha_k) = \int_0^1 \log \alpha(1 - \alpha) d\alpha. \quad (24)$$

795

806 ²An & Jeon (2023) enforce $D_j(\cdot, z)$ to be monotone in α (to avoid quantile crossing) via an isotonic-spline
807 parameterization. We do not impose this constraint: it adds architectural restrictions and, in our experiments,
808 occasional finite- K crossings had negligible effect on CRPS or downstream rollouts.

810 Hence $-\text{ELBO}_K$ converges to
 811

$$812 \mathcal{L}_{\text{DistVAE}}(\theta, \phi) = \mathbb{E}_{q_\phi(z|x)} \left[\sum_{j=1}^p \int_0^1 \rho_\alpha(x_j - D_j(\alpha, z; \theta_j)) d\alpha \right] + \beta \text{KL}(q_\phi(z|x) \| p(z)) + C, \\ 813 814 815 \quad (25)$$

816 where $\int_0^1 \rho_\alpha(\cdot) d\alpha$ equals the *Continuous Ranked Probability Score* (CRPS) for the model CDF. In
 817 practice we estimate it by Monte Carlo over $\alpha \sim \text{Unif}(0, 1)$. Thus the “ALD NLL (MC-CRPS)”
 818 reconstruction is the $K \rightarrow \infty$ limit of a valid ELBO (not a heuristic).
 819

820 C.3 WHY THIS HELPS VS. GAUSSIAN DECODING 821

822 **Distributional capacity.** Gaussian decoders impose symmetry and typically homoscedastic noise,
 823 and in practice often compensate for mean misspecification by *inflating the predicted variance*,
 824 yielding over-dispersed (underconfident) forecasts. ALD/quantile decoding directly captures *skewness* and *heteroscedasticity* across α while preserving VAE advantages: (i) a likelihood-derived
 825 proper scoring rule (CRPS) for reconstruction, (ii) simple sampling via inverse transform ($u \sim$
 826 $\text{Unif}(0, 1)$ then $x_j = D_j(u, z)$), (iii) a tractable latent KL. By focusing the loss on quantile locations
 827 across α , the ALD/CRPS objective discourages variance inflation and typically yields sharper
 828 predictive distributions under non-Gaussian data.
 829

830 C.4 OUR OBJECTIVE (CONTINUOUS HEAD) IN DISTVAE FORM 831

832 Identifying $x \equiv \mathbf{y}$ (continuous outcomes), our training loss for the outcome head is
 833

$$834 \mathcal{L}_{\text{cont}} = \mathbb{E}_{q_\phi(z|\cdot)} \left[\frac{1}{K} \sum_{k=1}^K \sum_{j=1}^{d_y} \rho_{\alpha^{(k)}}(y_j - D_j(\alpha^{(k)}, z; \theta_j)) \right] + \beta \text{KL}(q_\phi(z|\cdot) \| p(z)), \quad \alpha^{(k)} \stackrel{\text{i.i.d.}}{\sim} \text{Unif}(0, 1). \\ 835 836 837 \quad (26)$$

838 This is exactly the *ALD NLL (MC-CRPS)* plus KL, i.e., the continuous-variable DistVAE objective
 839 specialized to our architecture (temporal and cross-outcome dependence are mediated by the latent
 840 path; the quantile head supplies the likelihood noise, analogous to a Gaussian decoder’s noise).
 841

842 D G-LATENT ARCHITECTURE: ENCODER, TEMPORAL CORE, AND 843 DECODER

844 **Scope.** This appendix specifies the *network architecture* of G-Latent: the history network f_ω , the
 845 temporal core (κ_ψ and GRU_γ), and the shared conditional VAE (E_ϕ, D_θ) reused at every relative
 846 step. Training objectives and identification assumptions are described elsewhere.
 847

848 D.1 NOTATION AND SHAPES 849

850 Let $\mathbf{X}_t \in \mathbb{R}^{d_x}$, $\mathbf{Y}_t \in \mathbb{R}^{d_y}$, and $\mathbf{L}_t = (\mathbf{Y}_t, \mathbf{X}_t) \in \mathbb{R}^{d_L}$ with $d_L = d_x + d_y$; treatments $\mathbf{A}_t \in \mathbb{R}^{d_a}$; and
 851 static covariates $\mathbf{V} \in \mathbb{R}^{d_v}$. The history network outputs $\mathbf{r}_t \in \mathbb{R}^{d_r}$. At relative step $t' \in \{1, \dots, \tau\}$,
 852 the latent is $\mathbf{z}_{t,t'} \in \mathbb{R}^{d_z}$, the temporal state is $\mathbf{s}_{t,t'} \in \mathbb{R}^{d_s}$, and the step context is $\mathbf{c}_{t,t'} \in \mathbb{R}^{d_c}$.
 853

854 D.2 HISTORY NETWORK f_ω 855

856 We use the multi-input transformer of Melnychuk et al. (2022) (full details in App. B). Briefly:
 857

- 858 • **Inputs.** Three factual streams up to anchor time t : $\bar{\mathbf{X}}_t$, $\bar{\mathbf{Y}}_t$, and left-shifted $\bar{\mathbf{A}}_{t-1}$
 859 (start-of-interval indexing), plus static \mathbf{V} . Each stream is linearly projected to the model
 860 width; \mathbf{V} is injected at every time step.
- 861 • **Blocks.** Each block applies masked multi-head self-attention with shared relative
 862 positional encodings per stream, cross-attentions between streams, and a positionwise
 863 feed-forward network. All sublayers use residual connections, layer normalization, and
 864 dropout.

864 • **Fusion.** The final per-time states of the three streams are averaged and linearly projected
 865 with ELU to yield $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t) \in \mathbb{R}^{d_r}$. No model predictions are fed back into the encoder.
 866

867 D.3 TEMPORAL CORE: CONTEXT COMBINER AND LATENT-DRIVEN STATE UPDATE
 868

869 Given \mathbf{r}_t , previous state $\mathbf{s}_{t,t'-1}$, current action $\mathbf{a}_{t+t'-1}$, and relative index t' , we form a dense context
 870 and update the recurrent state.

871 **Context combiner.** We concatenate the inputs and project to d_c with a single linear layer:

872 $\tilde{\mathbf{c}}_{t,t'} = [\mathbf{r}_t ; \mathbf{s}_{t,t'-1} ; \mathbf{a}_{t+t'-1} ; t'] \in \mathbb{R}^{d_r+d_s+d_a+1}, \quad \mathbf{c}_{t,t'} = \kappa_\psi(\tilde{\mathbf{c}}_{t,t'}) \in \mathbb{R}^{d_c}. \quad (27)$
 873

875 **State update (latents only).** A GRUCell updates the temporal state using the latent, the frozen
 876 history embedding, the current action, and the step index:

877 $\mathbf{s}_{t,t'} = \text{GRU}_\gamma([\mathbf{z}_{t,t'} ; \mathbf{r}_t ; \mathbf{a}_{t+t'-1} ; t'], \mathbf{s}_{t,t'-1}), \quad \mathbf{s}_{t,0} = \mathbf{0}. \quad (28)$
 878

879 GRU weights are orthogonally initialized and biases are zero-initialized. A data-space variant (not
 880 used in our main model) replaces $\mathbf{z}_{t,t'}$ with $\mathbf{l}_{t+t'}$.

881 D.4 SHARED CONDITIONAL VAE (E_ϕ, D_θ)
 882

883 A single conditional VAE is reused across steps. Encoder E_ϕ outputs a Gaussian posterior over $\mathbf{z}_{t,t'}$,
 884 and decoder D_θ maps $[\mathbf{z}_{t,t'}; \mathbf{c}_{t,t'}; \mathbf{a}_{t+t'-1}]$ to the reconstruction heads. The decoder uses *dense skip*
 885 concatenation: after every hidden block, $[\mathbf{z}; \mathbf{c}; \mathbf{a}]$ is re-concatenated to the block output before the
 886 next block.

887 D.4.1 ENCODER E_ϕ
 888

889 The encoder is an MLP applied to $[\mathbf{l}_{t+t'}; \mathbf{c}_{t,t'}]$ with repeated blocks Linear \rightarrow BatchNorm \rightarrow
 890 ReLU \rightarrow Dropout, followed by two linear heads for mean and log-variance:

891 $(\boldsymbol{\mu}_{t,t'}, \log \boldsymbol{\sigma}_{t,t'}^2) = E_\phi([\mathbf{l}_{t+t'}; \mathbf{c}_{t,t'}]) \in \mathbb{R}^{d_z} \times \mathbb{R}^{d_z}, \quad \mathbf{z}_{t,t'} = \boldsymbol{\mu}_{t,t'} + \boldsymbol{\sigma}_{t,t'} \odot \boldsymbol{\epsilon}, \quad \boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}). \quad (29)$
 892

893 D.4.2 DECODER TRUNK T_θ WITH DENSE SKIPS
 894

895 Starting from $\mathbf{h}_0 = [\mathbf{z}_{t,t'}; \mathbf{c}_{t,t'}; \mathbf{a}_{t+t'-1}]$, the trunk applies repeated blocks Linear \rightarrow ReLU \rightarrow
 896 Dropout; after each block with output \mathbf{h} , we set

897 $\mathbf{h} \leftarrow [\mathbf{h} ; \mathbf{z}_{t,t'} ; \mathbf{c}_{t,t'} ; \mathbf{a}_{t+t'-1}] \quad (30)$
 898

899 before entering the next block. The trunk output $\mathbf{w}_{t,t'}$ feeds the heads below.

900 **Gaussian (heteroscedastic) decoding path.** When using a purely Gaussian decoder for all d_L
 901 coordinates, two linear heads produce mean and positive scale (via softplus):

902 $\hat{\boldsymbol{\mu}}_{t,t'} = W_\mu \mathbf{w}_{t,t'} + b_\mu, \quad \hat{\boldsymbol{\sigma}}_{t,t'} = \text{softplus}(W_\sigma \mathbf{w}_{t,t'} + b_\sigma), \quad (31)$
 903

904 yielding a diagonal Gaussian on $\mathbf{L}_{t+t'}$. Optional clamping can be applied to designated coordinates
 905 (e.g., nonnegativity of specific outputs) by shifting the corresponding mean channels.

906 **CRPS / random-quantile outcome path.** When using the distributional outcome head, the de-
 907 coder splits into:

1. **Outcome quantile head (per outcome, per quantile).** Let $\boldsymbol{\alpha} \in (0, 1)^{d_y}$ collect
 910 per-outcome quantile levels and draw A i.i.d. samples per outcome. From $\mathbf{w}_{t,t'}$ (option-
 911 ally after a small shared sub-trunk), each outcome coordinate $j \in \{1, \dots, d_y\}$ has a dedi-
 912 cated MLP that *re-concatenates* $[\mathbf{z}_{t,t'}; \mathbf{c}_{t,t'}; \mathbf{a}_{t+t'-1}; \alpha_j]$ at every hidden layer and outputs
 913 a scalar quantile $\hat{q}_{\alpha_j, t, t', j}$. Stacking across A samples yields $\hat{\mathbf{Q}}_{t,t'} \in \mathbb{R}^{d_y \times A}$.
2. **Remaining coordinates (Gaussian head).** If $d_L > d_y$, a separate trunk (fed by
 915 $[\mathbf{w}_{t,t'}; \mathbf{z}_{t,t'}; \mathbf{c}_{t,t'}; \mathbf{a}_{t+t'-1}]$) outputs $(\hat{\boldsymbol{\mu}}_{\text{rem}}, \log \hat{\boldsymbol{\sigma}}_{\text{rem}}^2)$ for the remaining $d_L - d_y$ coordinates.

917 This realizes the outcome-specific α -aware branches while keeping non-outcome channels Gaus-
 918 sian.

918 D.5 PER-STEP FLOW (TRAINING AND INFERENCE INTERFACE)
919920 At each step t' :

921 1. Build the context:

922
$$\mathbf{c}_{t,t'} = \kappa_\psi([\mathbf{r}_t; \mathbf{s}_{t,t'-1}; \mathbf{a}_{t+t'-1}; t']). \quad (32)$$

924 2. *Training*: encode $[\mathbf{l}_{t+t'}; \mathbf{c}_{t,t'}]$ to obtain $(\boldsymbol{\mu}_{t,t'}, \log \boldsymbol{\sigma}_{t,t'}^2)$ and sample $\mathbf{z}_{t,t'}$.925 3. Decode with either the Gaussian head to obtain $(\hat{\boldsymbol{\mu}}_{t,t'}, \hat{\boldsymbol{\sigma}}_{t,t'})$ for all coordinates, or the
926 quantile outcome head to obtain $\hat{\mathbf{Q}}_{t,t'}$ (and Gaussian parameters for any remaining coordi-
927 nates).

928 4. Update the state:

929
$$\mathbf{s}_{t,t'} = \text{GRU}_\gamma([\mathbf{z}_{t,t'}; \mathbf{r}_t; \mathbf{a}_{t+t'-1}; t'], \mathbf{s}_{t,t'-1}). \quad (33)$$

930 At inference, $\mathbf{z}_{t,t'} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ is sampled independently across steps and Monte Carlo draws; by
931 default only outcomes \mathbf{Y} are decoded, and decoding can be restricted to any subset of steps $S \subseteq$
932 $\{1, \dots, \tau\}$.
933935 D.6 DESIGN NOTES
936

- **Treatment sensitivity.** Actions enter both the context combiner and *every* decoder block via dense re-concatenation, preserving a short path from treatment to outputs.
- **Relative step embedding.** The scalar index t' (or a small positional code) is concatenated in κ_ψ and the GRU input to inform the horizon position without per-step parameters.
- **Normalization and positivity.** BatchNorm is used only in the VAE encoder. Decoder scales are enforced positive with softplus.
- **Parameter sharing.** A single $(E_\phi, D_\theta, \kappa_\psi, \text{GRU}_\gamma)$ instance is reused across all t' , improving data efficiency and keeping semantics consistent across horizons.

946 D.7 MODULE I/O SUMMARY
947

Module	Signature
History network f_ω	$\bar{\mathbf{h}}_t \mapsto \mathbf{r}_t \in \mathbb{R}^{d_r}$
Context combiner κ_ψ	$[\mathbf{r}_t; \mathbf{s}_{t,t'-1}; \mathbf{a}_{t+t'-1}; t'] \mapsto \mathbf{c}_{t,t'} \in \mathbb{R}^{d_c}$
Encoder E_ϕ	$[\mathbf{l}_{t+t'}; \mathbf{c}_{t,t'}] \mapsto (\boldsymbol{\mu}_{t,t'}, \log \boldsymbol{\sigma}_{t,t'}^2) \in \mathbb{R}^{d_z} \times \mathbb{R}^{d_z}$
Decoder trunk T_θ	$[\mathbf{z}_{t,t'}; \mathbf{c}_{t,t'}; \mathbf{a}_{t+t'-1}] \mapsto \mathbf{w}_{t,t'} \text{ (dense skips)}$
Outcome head $D_\theta^{(y)}$	$(\text{CRPS}) [\mathbf{w}_{t,t'}; \alpha] \mapsto \hat{q}_\alpha \in \mathbb{R} \text{ (per outcome, per } \alpha)$
Covariate head $D_\theta^{(x)}$	$(\text{Gaussian}) \mathbf{w}_{t,t'} \mapsto (\hat{\boldsymbol{\mu}}_{\text{rem}}, \log \hat{\boldsymbol{\sigma}}_{\text{rem}}^2)$
State update GRU_γ	$[\mathbf{z}_{t,t'}; \mathbf{r}_t; \mathbf{a}_{t+t'-1}; t'], \mathbf{s}_{t,t'-1} \mapsto \mathbf{s}_{t,t'}$

958 E THEORETICAL INSIGHTS
959960 E.1 EQUIVALENCE OF LATENT AND DATA-SPACE G-COMPUTATION
961962 We first formalize when sampling *only in latent space* (Alg. 1) is sufficient to recover the interven-
963 tional laws identified by the sequential g-formula.
964965 **Standing causal assumptions.** We assume the usual conditions for identification by the g-
966 formula: (i) consistency, (ii) sequential ignorability/exchangeability, and (iii) sequential over-
967 lap/positivity (cf. App. A).968 **Assumption E.1** (Latent factorization and context sufficiency). Let p_0 be a fixed prior density on
969 latents (e.g., $\mathcal{N}(\mathbf{0}, \mathbf{I})$). Fix an anchor time t and let $\mathbf{r}_t = f_\omega(\bar{\mathbf{h}}_t)$ denote the history embedding
970 computed at t . For each relative step $t' \in \{1, \dots, \tau\}$ define the latent-state update recursively from
971 $\mathbf{s}_{t,0} = \mathbf{0}$ by

972
$$\mathbf{s}_{t,t'} = \text{GRU}_\gamma([\mathbf{z}_{t,t'}, \mathbf{r}_t, \mathbf{a}_{t+t'-1}, t'], \mathbf{s}_{t,t'-1}).$$

972 Assume that for every $t' \in \{1, \dots, \tau\}$ and every history $\bar{\mathbf{h}}_{t+t'-1}$ the true one-step conditional
 973 distribution of $\mathbf{L}_{t+t'} = (\mathbf{Y}_{t+t'}, \mathbf{X}_{t+t'})$ admits the factorization
 974

$$975 \quad p^*(\mathbf{l}_{t+t'} \mid \bar{\mathbf{h}}_{t+t'-1}, \mathbf{a}_{t+t'-1}) = \int p_\theta(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}) p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'},$$

977 where $\mathbf{c}_{t,t'} = \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$. Moreover, for each fixed (\mathbf{c}, \mathbf{a}) the map $\mathbf{z} \mapsto p_\theta(\cdot \mid$
 978 $\mathbf{z}, \mathbf{c}, \mathbf{a})$ is a probability-kernel in \mathbf{L} measurable in $(\mathbf{z}, \mathbf{c}, \mathbf{a})$.
 979

980 Assumption E.1 states that $(\mathbf{r}_t, \mathbf{s}_{t,t'-1})$ is a *sufficient statistic* of $\bar{\mathbf{H}}_{t+t'-1}$ for predicting $\mathbf{L}_{t+t'}$, and
 981 that the true stepwise conditional factors through a latent with fixed prior density p_0 .
 982

983 *Remark E.2* (Relation to training). Assumption E.1 is a modeling/realizability statement: it pos-
 984 tulates that the one-step conditionals factor through a latent with prior p_0 given the context
 985 $(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$. Our conditional-VAE training (Sec. 4.3) is the estimation procedure we use
 986 to realize this factorization in practice by maximizing the (conditional) ELBO, i.e., approximately
 987 minimizing the negative log-likelihood of $p_\theta(\mathbf{L}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$ under p_0 . All results that
 988 require the assumption hold exactly; with finite data and imperfect training, they hold approximately
 989 with the local errors $\{\varepsilon_{t'}\}$ used in Prop. E.10.

989 *Remark E.3* (State update uses latent representations). The recurrent state is updated *through latents*
 990 *only* ($\mathbf{s}_{t,t'} = \text{GRU}_\gamma([\mathbf{z}_{t,t'}, \mathbf{r}_t, \mathbf{a}_{t+t'-1}, t'], \mathbf{s}_{t,t'-1})$). Thus all predictive information that propagates
 991 forward from step t' enters via $\mathbf{z}_{t,t'}$ and the context $\mathbf{c}_{t,t'} = \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$. When $\mathbf{z}_{t,t'}$ is
 992 a *good representation* of $\mathbf{L}_{t+t'}$ (e.g., the decoder $p_\theta(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$ is highly expressive
 993 and, ideally, injective in $\mathbf{z}_{t,t'}$ for a.e. $(\mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$), the pair $(\mathbf{r}_t, \mathbf{s}_{t,t'-1})$ approaches a sufficient
 994 statistic of $\bar{\mathbf{H}}_{t+t'-1}$ for predicting $\mathbf{L}_{t+t'}$. In VAEs the mapping is not exactly invertible, but training
 995 to maximize the conditional ELBO encourages $\mathbf{z}_{t,t'}$ to retain information about $\mathbf{L}_{t+t'}$ that is relevant
 996 for prediction; higher-fidelity decoders (e.g., with flows) make this approximation tighter.
 997

998 **Lemma E.4** (Representation sufficiency implies context sufficiency). *Fix the embedding \mathbf{r}_t and*
 999 *suppose Assumption E.1 holds. Assume that for Lebesgue-a.e. (\mathbf{c}, \mathbf{a}) the mapping $\mathbf{z} \mapsto p_\theta(\cdot \mid \mathbf{z}, \mathbf{c}, \mathbf{a})$*
 1000 *is injective as a map into $\mathcal{P}(\mathcal{L})$ (i.e., distinct \mathbf{z} induce distinct conditional laws). Assume also that*
 1001 *$\mathbf{s}_{t,t'-1}$ is a deterministic, measurable function of $(\mathbf{z}_{t,1:t'-1}, \mathbf{r}_t, \mathbf{a}_{t:t'-2}, 1:t'-1)$. Then for almost*
 1002 *every $(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$ we have the conditional independence*

$$1002 \quad \mathbf{L}_{t+t'} \perp\!\!\!\perp \bar{\mathbf{H}}_{t+t'-1} \mid (\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t'),$$

1003 *i.e., $(\mathbf{r}_t, \mathbf{s}_{t,t'-1})$ is sufficient (with $\mathbf{a}_{t+t'-1}, t'$) for predicting $\mathbf{L}_{t+t'}$.*
 1004

1005 *Proof.* Given $(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$, the next context $\mathbf{c}_{t,t'}$ is fixed and $\mathbf{z}_{t,t'} \sim p_0$ is independent
 1006 of $\bar{\mathbf{H}}_{t+t'-1}$. The conditional density of $\mathbf{L}_{t+t'}$ factors as $p(\mathbf{l}_{t+t'} \mid \bar{\mathbf{h}}_{t+t'-1}, \mathbf{a}_{t+t'-1}) = \int p_\theta(\mathbf{l}_{t+t'} \mid$
 1007 $\mathbf{z}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}) p_0(\mathbf{z}) d\mathbf{z}$ by Assumption E.1. Because $\mathbf{s}_{t,t'-1}$ is a deterministic function of past
 1008 latents, any dependence on $\bar{\mathbf{H}}_{t+t'-1}$ enters only through $(\mathbf{r}_t, \mathbf{s}_{t,t'-1})$. Injectivity in \mathbf{z} rules out
 1009 aliasing of predictive distributions conditioned on $\mathbf{c}_{t,t'}$, so conditioning on $(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$
 1010 screens off the past. \square
 1011

1012 **Notation.** When we write $\mathbf{c}_{t,t'}(\mathbf{z}_{t,1:t'-1})$ we suppress fixed arguments $(\mathbf{r}_t, \mathbf{a}_{t+t'-1}, t')$ and em-
 1013 phasize the indirect dependence via $\mathbf{s}_{t,t'-1}$; explicitly, $\mathbf{c}_{t,t'} = \kappa_\psi(\mathbf{r}_t, \mathbf{s}_{t,t'-1}(\mathbf{z}_{t,1:t'-1}), \mathbf{a}_{t+t'-1}, t')$.
 1014

1015 **Theorem E.5** (Equivalence of latent and data-space g-computation). *Fix a time t , a horizon $\tau \geq 1$,*
 1016 *a treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$, and a history $\bar{\mathbf{h}}_t$. Under the standing causal assumptions and Assump-*
 1017 *tion E.1, the interventional law identified by the sequential g-formula equals the law induced by*
 1018 *latent rollout (Alg. 1):*
 1019

1020 (i) **(Fixed-horizon marginal)** For the last-step outcome,

$$1022 \quad p^{\bar{\mathbf{a}}}(\mathbf{y}_{t+\tau} \mid \bar{\mathbf{h}}_t) = \int p_\theta(\mathbf{y}_{t+\tau} \mid \mathbf{z}_{t,\tau}, \mathbf{c}_{t,\tau}(\mathbf{z}_{t,1:\tau-1}), \mathbf{a}_{t+\tau-1}) \prod_{t'=1}^{\tau} p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'}.$$

1023 Here $\mathbf{c}_{t,\tau}(\mathbf{z}_{t,1:\tau-1})$ is the deterministic context produced by the latent-state recursion
 1024 driven by $\mathbf{z}_{t,1:\tau-1}$.
 1025

1026 (ii) (**Full-path law**) For the joint path,

$$1028 \quad p^{\bar{a}}(\mathbf{y}_{t+1:t+\tau} \mid \bar{\mathbf{h}}_t) = \int \prod_{t'=1}^{\tau} p_{\theta}(\mathbf{y}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}(\mathbf{z}_{t,1:t'-1}), \mathbf{a}_{t+t'-1}) \prod_{t'=1}^{\tau} p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'},$$

1030 where, if desired, the covariates $\{\mathbf{x}_{t+t'}\}$ are integrated out.

1031 Consequently, the Monte Carlo samples produced by Alg. 1 (with `scope=last` or `all`) are i.i.d.
1032 draws from the respective interventional laws.

1033 *Proof.* By identification, the last-step interventional density is

$$1037 \quad p^{\bar{a}}(\mathbf{y}_{t+\tau} \mid \bar{\mathbf{h}}_t) = \int_{\mathbf{l}_{t+1:t+\tau-1}} \left[\prod_{t'=1}^{\tau-1} p(\mathbf{l}_{t+t'} \mid \bar{\mathbf{h}}_{t+t'-1}, \mathbf{a}_{t+t'-1}) \right] p(\mathbf{y}_{t+\tau} \mid \bar{\mathbf{h}}_{t+\tau-1}, \mathbf{a}_{t+\tau-1}) d\mathbf{l}_{t+1:t+\tau-1}.$$

1040 Insert Assumption E.1 at each step (including the last) to obtain

$$1042 \quad \int \left\{ \int_{\mathbf{l}_{t+1:t+\tau-1}} \prod_{t'=1}^{\tau-1} p_{\theta}(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1}) d\mathbf{l}_{t+1:t+\tau-1} \right\} p_{\theta}(\mathbf{y}_{t+\tau} \mid \mathbf{z}_{t,\tau}, \mathbf{c}_{t,\tau}, \mathbf{a}_{t+\tau-1}) \prod_{t'=1}^{\tau} p_0(\mathbf{z}_{t,t'}) d\mathbf{z}_{t,t'}.$$

1045 Using Tonelli/Fubini (all integrands are nonnegative densities), we can swap integration order, and
1046 since $p_{\theta}(\mathbf{l}_{t+t'} \mid \mathbf{z}_{t,t'}, \mathbf{c}_{t,t'}, \mathbf{a}_{t+t'-1})$ is a normalized conditional density with $\mathbf{c}_{t,t'}$ independent of
1047 decoded \mathbf{L} , we have $\int p_{\theta}(\mathbf{l}_{t+t'} \mid \cdot) d\mathbf{l}_{t+t'} = 1$ for $t' = 1, \dots, \tau - 1$ (the remaining integrals are
1048 over the latent path $\mathbf{z}_{t:t+\tau}$ and the terminal outcome, i.e., we are integrating out all intermediate
1049 variables). This yields the first result (i).

1050 For the full-path law, repeat the same steps but keep the outcome components $\mathbf{y}_{t+t'}$ unintegrated
1051 (integrate only the covariates $\mathbf{x}_{t+t'}$ if desired). The product form in item (ii) follows because $\mathbf{c}_{t,t'}$
1052 depends only on $(\mathbf{r}_t, \mathbf{a}_{t:t+t'-1}, \mathbf{z}_{t,1:t'-1})$, never on decoded \mathbf{L} . Finally, Alg. 1 draws $\{\mathbf{z}_{t,t'}\}$ i.i.d.
1053 from p_0 and applies the same deterministic maps and decoder conditional densities as above, so its
1054 outputs are i.i.d. from these laws. \square

1055 **Corollary E.6** (Selective decoding (`scope`) is coherent). *Under the conditions of Thm. E.5, de-
1056 coding only at $t+\tau$ (`scope=last`) returns i.i.d. samples from $p^{\bar{a}}(\mathbf{y}_{t+\tau} \mid \bar{\mathbf{h}}_t)$. More generally,
1057 decoding at any subset $S \subseteq \{1, \dots, \tau\}$ returns i.i.d. samples from the corresponding marginal over
1058 $\{\mathbf{Y}_{t+t'}\}_{t' \in S}$.*

1059 **Sketch / intuition.** The sequential g-formula integrates over *future observations*. Assumption E.1
1060 lets each one-step conditional be written as a mixture over a *latent noise* $\mathbf{z}_{t,t'}$ whose context depends
1061 only on $(\mathbf{r}_t, \mathbf{s}_{t,t'-1}, \mathbf{a}_{t+t'-1}, t')$. Because future contexts never use decoded \mathbf{L} , all intermediate
1062 integrals over \mathbf{L} collapse to 1: only the latent-driven contexts matter. Thus, sampling latents and
1063 decoding where desired reproduces the same interventional law.

1064

1066 E.2 PROPAGATION ERROR: LATENT VS. DATA-SPACE G-COMPUTATION ROLLOUT

1067 We now provide theoretical justification for the empirical superiority of latent rollouts over autore-
1068 gressive rollouts. Let $\{K_s^*\}_{s=t}^{t+\tau-1}$ denote the true one-step transition kernels and $\{K_s^e\}_{s=t}^{t+\tau-1}$ the
1069 learned approximations. For each step s , define the local one-step approximation error

$$1071 \quad \varepsilon_s := \sup_{\bar{h}_s, a_s} \text{TV}\left(K_s^*(\cdot \mid \bar{h}_s, a_s), K_s^e(\cdot \mid \bar{h}_s, a_s)\right),$$

1073 where TV denotes total variation distance.

1075 **Tail operators.** For $s \in \{t, \dots, t+\tau-1\}$, let $T_{s+1:t+\tau}^{\bullet}$ denote the *tail operator* that maps a law on
1076 L_{s+1} to the induced law of $Y_{t+\tau}$ obtained by propagating forward under rollout type $\bullet \in \{\text{lat}, \text{AR}\}$.

1077 It is standard that pushing forward measures by a fixed Markov kernel is nonexpansive in total
1078 variation, hence

$$1079 \quad \text{TV}(T_{s+1:t+\tau}^{\text{lat}}[\mu], T_{s+1:t+\tau}^{\text{lat}}[\nu]) \leq \text{TV}(\mu, \nu). \quad (34)$$

1080 The property is a standard result for Markov kernels, often referred to as the Data Processing In-
 1081 equality. For autoregressive rollouts, the re-encoding step introduces sensitivity to the input measure.
 1082 We assume:

1083 **Assumption E.7** (Single-step AR operator Lipschitz property). For each index $j \in \{t+1, \dots, t+\tau-1\}$ define the *single-step AR tail operator*

$$1086 \mathcal{T}_{j \rightarrow j+1}^{\text{AR}} : \mathcal{P}(\mathcal{L}_j) \longrightarrow \mathcal{P}(\mathcal{L}_{j+1}),$$

1087 which maps a law on L_j (the predicted/decoded quantity at time j) to the induced law of the next-
 1088 step quantity under the autoregressive re-encoding and decoding procedure. Assume there exist
 1089 constants $\lambda_j \geq 0$ such that, for all probability measures μ, ν on \mathcal{L}_j ,

$$1091 \text{TV}(\mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\mu], \mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\nu]) \leq (1 + \lambda_j) \text{TV}(\mu, \nu).$$

1093 The assumption E.7 is justified because the autoregressive operator, as a finite composition of linear
 1094 layers and Lipschitz-continuous activation functions, is itself guaranteed to be Lipschitz-continuous
 1095 on any bounded domain.

1096 **Lemma E.8** (Composition amplification). *Under Assumption E.7, the composed AR tail operator*
 1097 $T_{s+1:t+\tau}^{\text{AR}} = \mathcal{T}_{t+\tau-1 \rightarrow t+\tau}^{\text{AR}} \circ \dots \circ \mathcal{T}_{s+1 \rightarrow s+2}^{\text{AR}}$ *satisfies, for any μ, ν on \mathcal{L}_{s+1} ,*

$$1099 \text{TV}(T_{s+1:t+\tau}^{\text{AR}}[\mu], T_{s+1:t+\tau}^{\text{AR}}[\nu]) \leq \prod_{j=s+1}^{t+\tau-1} (1 + \lambda_j) \text{TV}(\mu, \nu).$$

1102 *Proof.* Apply the single-step bound (S) iteratively. For brevity write $\mu_{s+1} = \mu$, $\nu_{s+1} = \nu$ and
 1103 define $\mu_{j+1} = \mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\mu_j]$, $\nu_{j+1} = \mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\nu_j]$. Then

$$1105 \text{TV}(\mu_{j+1}, \nu_{j+1}) \leq (1 + \lambda_j) \text{TV}(\mu_j, \nu_j).$$

1106 Chaining these inequalities for $j = s+1, \dots, t+\tau-1$ yields

$$1108 \text{TV}(\mu_{t+\tau}, \nu_{t+\tau}) \leq \left(\prod_{j=s+1}^{t+\tau-1} (1 + \lambda_j) \right) \text{TV}(\mu_{s+1}, \nu_{s+1}),$$

1111 which is the claimed bound. □

1113 **Remark E.9** (A sufficient bound for λ_j). A convenient sufficient condition for Assumption E.7 is
 1114 obtained by decomposing the single-step AR operator into (i) a *re-encoding map* $\Xi_j : \mathcal{P}(\mathcal{L}_j) \rightarrow$
 1115 \mathcal{C}_j that maps a predicted law on L_j to a context in \mathcal{C}_j , and (ii) a decoder-induced kernel family
 1116 $\{K_j^c\}_{c \in \mathcal{C}_j}$ that maps a context to a next-step kernel.

1117 Concretely, suppose that for each j :

1118 1. Ξ_j is $L_{\Xi,j}$ -Lipschitz in total variation, i.e.

$$1120 \text{TV}(\Xi_j[\mu], \Xi_j[\nu]) \leq L_{\Xi,j} \text{TV}(\mu, \nu) \quad \text{for all } \mu, \nu \in \mathcal{P}(\mathcal{L}_j);$$

1122 2. the decoder-induced kernel family is $L_{K,j}$ -Lipschitz in context, i.e.

$$1124 \sup_{c, c'} \text{TV}(K_j^c, K_j^{c'}) \leq L_{K,j} \|c - c'\|.$$

1126 Then for any two input measures μ, ν on \mathcal{L}_j we have

$$1128 \text{TV}(\mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\mu], \mathcal{T}_{j \rightarrow j+1}^{\text{AR}}[\nu]) \leq \sup_{c, c'} \text{TV}(K_j^c, K_j^{c'}) \leq L_{K,j} \|\Xi_j[\mu] - \Xi_j[\nu]\| \leq L_{K,j} L_{\Xi,j} \text{TV}(\mu, \nu).$$

1130 Hence one may take

$$1132 \lambda_j \leq L_{K,j} L_{\Xi,j},$$

1133 and the product amplification in Proposition E.10 follows by composing these single-step bounds
 (cf. Lemma E.8).

1134 **Main result.** Let P^* denote the true marginal law of $Y_{t+\tau}$, P^{lat} the law induced by the latent
 1135 rollout, and P^{AR} the law induced by the autoregressive rollout. Then:

1136 **Proposition E.10** (Propagation-error bound and dominance). *Let $t, \tau, a_{t:t+\tau-1}, \bar{h}_t$ be fixed, and let
 1137 P^* denote the true interventional law of $Y_{t+\tau}$. Let P^{lat} and P^{AR} denote the learned laws produced
 1138 by the latent and autoregressive/data-space rollouts, respectively, when both use the same per-step
 1139 approximations $\{K_s^e\}_{s=t}^{t+\tau-1}$. Under Assumption E.7 (single-step AR operator Lipschitz property)
 1140 we have*

$$1141 \quad \text{TV}(P^*, P^{\text{lat}}) \leq \sum_{s=t}^{t+\tau-1} \varepsilon_s, \quad (35)$$

$$1145 \quad \text{TV}(P^*, P^{\text{AR}}) \leq \sum_{s=t}^{t+\tau-1} \varepsilon_s \prod_{j=s+1}^{t+\tau-1} (1 + \lambda_j), \quad (36)$$

1148 where $\varepsilon_s := \sup_{\bar{h}_s, a_s} \text{TV}(K_s^*(\cdot \mid \bar{h}_s, a_s), K_s^e(\cdot \mid \bar{h}_s, a_s))$. In particular, if some $\lambda_j > 0$ then the
 1149 bound equation 36 dominates equation 35, so the latent rollout attains a uniformly tighter (or equal)
 1150 upper bound on the final-step discrepancy.

1152 *Proof.* For the latent rollout, a standard telescoping decomposition across steps combined with the
 1153 non-expansive property of Markov kernels in Equation 34 yields the bound:

$$1155 \quad \text{TV}(P^*, P^{\text{lat}}) \leq \sum_{s=t}^{t+\tau-1} \varepsilon_s.$$

1158 For the autoregressive rollout, we define a sequence of hybrid distributions P_s for $s = t, \dots, t + \tau$,
 1159 where P_s is the law generated by using the true kernels K^* up to step $s - 1$ and the learned kernels
 1160 K^e from step s onwards. This gives $P_{t+\tau} = P^*$ and $P_t = P^{\text{AR}}$.

1161 By the triangle inequality, the total error is bounded by the sum of one-step differences:

$$1163 \quad \text{TV}(P^*, P^{\text{AR}}) = \text{TV}(P_{t+\tau}, P_t) \leq \sum_{s=t}^{t+\tau-1} \text{TV}(P_{s+1}, P_s).$$

1166 The difference between P_{s+1} and P_s arises only from the kernel used at step s . The error introduced
 1167 at this step, at most ε_s , is then propagated forward by the autoregressive tail operator $T_{s+1:t+\tau}^{\text{AR}}$.
 1168 Using the amplification bound from Lemma E.8, the contribution from step s is:

$$1170 \quad \text{TV}(P_{s+1}, P_s) \leq \varepsilon_s \prod_{j=s+1}^{t+\tau-1} (1 + \lambda_j).$$

1173 Summing these terms from $s = t$ to $t + \tau - 1$ yields the bound in Equation 36. Since each factor
 1174 $(1 + \lambda_j) \geq 1$, the bound in Equation 36 is uniformly greater than or equal to the bound in Equation
 1175 35, completing the proof. \square

F DATASETS

F.1 DETAILS ON EXPERIMENTS WITH SEMI-SYNTHETIC DATA (ORIGINAL SETTING)

1181 Following Melnychuk et al. (2022), we build on MIMIC-EXTRACT (Wang et al., 2020)—a stan-
 1182 dardized preprocessing pipeline for MIMIC-III (Johnson et al., 2016)—which provides ICU time
 1183 series aggregated at an hourly cadence. Missing values are imputed using forward and backward
 1184 filling, and all continuous time-varying variables are standardized.

1185 From this resource we retain 25 vital signs as time-varying covariates and three static covariates
 1186 (gender, ethnicity, age). The complete feature list is provided in the accompanying code repository
 1187 for reproducibility. Static covariates are one-hot encoded and later reused to modulate noise terms.
 1188 In total, this yields a $d_v = 44$ dimensional covariate vector.

1188
 1189 **High-level simulator design.** Following the basic idea of Schulam & Saria (2017), we first syn-
 1190 thesize *untreated* outcome trajectories under endogenous and exogenous dependencies, and then
 1191 apply treatments sequentially. We assume sparsity: each outcome depends on only a small subset
 1192 of covariates and treatments; treatment assignment likewise depends on a limited subset of recent
 1193 outcomes and covariates.

1194 **Cohort selection.** We sample 1,000 patients whose ICU stays last at least 20 hours. Stays longer
 1195 than 100 hours are clipped, so for patient i we have $T^{(i)} \in [20, 100]$.

1197 **Untreated outcomes.** For each patient i and each outcome dimension $j = 1, \dots, d_y$, we construct
 1198 an untreated signal $Z_{j,t}^{(i)}$ by combining (i) a global trend, (ii) a patient-specific smooth component,
 1199 (iii) an exogenous effect of current covariates, and (iv) noise:

1201
$$Z_{j,t}^{(i)} = \underbrace{\alpha_j^S \text{B-spline}(t)}_{\text{endogenous}} + \underbrace{\alpha_j^g g_j^{(i)}(t)}_{\text{exogenous}} + \underbrace{\varepsilon_t}_{\text{noise}}, \quad \varepsilon_t \sim \mathcal{N}(0, 0.005^2). \quad (37)$$

1204 Here, B-spline(t) is drawn from a mixture of three cubic splines (rapid decline, mild decline, stable)
 1205 over the ICU stay; $g_j^{(i)}(t)$ is an independent Gaussian process with a Matérn kernel; and $f_j^Z(\cdot)$ is
 1206 sampled via a random Fourier features (RFF) approximation to a Gaussian process (?), which avoids
 1207 repeated Cholesky factorizations when sampling at many points in \mathbb{R}^{d_x} . The weights $\alpha_j^S, \alpha_j^g, \alpha_j^f$
 1208 control the relative contributions.

1210 **Treatment assignment.** We then generate d_a binary treatments $\{A_t^l\}_{l=1}^{d_a}$ sequentially, introduc-
 1211 ing confounding through (a) a function of current covariates and (b) recent outcome history. For
 1212 treatment l at time t we define

1214
$$p_{l,t}^A = \sigma(\gamma_l^A \bar{A}_{T_l}(\bar{Y}_{t-1}) + \gamma_l^X f_l^Y(X_t) + b_l), \quad (38)$$

1216
$$A_t^l \sim \text{Bernoulli}(p_{l,t}^A), \quad (39)$$

1218 where $\sigma(\cdot)$ is the logistic function; $\bar{A}_{T_l}(\bar{Y}_{t-1})$ denotes the average over a selected subset of the
 1219 previous T_l treated outcomes using the history \bar{Y}_{t-1} ; $f_l^Y(\cdot)$ is sampled via an RFF GP (analogous
 1220 to f_j^Z); and γ_l^A, γ_l^X together with bias b_l govern the strength of confounding.

1222 **Treatment effects.** We set $Y_{j,1} = Z_{j,1}$ and endow each treatment l with a long-lasting additive
 1223 effect on outcome j that is maximal immediately after administration and decays as an inverse square
 1224 of elapsed time within a window of length w_l . Effects are scaled by the assignment probability $p_{l,i}^A$.
 1225 When multiple treatments are active, we aggregate their contributions conservatively by taking the
 1226 minimum at each elapsed time. Let $\varphi_l(\Delta) = (\Delta + 1)^{-2} \mathbf{1}\{0 \leq \Delta \leq w_l\}$. Then

1228
$$E_j(t) = \sum_{i=1}^t \min_{l=1, \dots, d_a} \left\{ \mathbf{1}\{A_i^l = 1\} p_{l,i}^A \beta_{lj} \varphi_l(t - i) \right\}, \quad (40)$$

1231 where β_{lj} is the maximum (immediate) effect size of treatment l on outcome j (either a constant or
 1232 zero if treatment l does not act on j).

1234 **Observed outcomes.** The observed process adds treatment effects to the untreated signal:

1236
$$Y_{j,t} = Z_{j,t} + E_j(t). \quad (41)$$

1238 **Dataset construction and evaluation.** Unless stated otherwise, exact simulator hyperparameters
 1239 are provided in the code. In our main setting we use $d_a = 3$ synthetic binary treatments and $d_y = 2$
 1240 outcomes. The 1,000 patients are split into train/validation/test using a 60%/20%/20% split. For
 1241 one-step-ahead evaluation we enumerate all $2^3 = 8$ counterfactuals. For multi-step rollouts with
 $\tau_{\max} = 10$, we sample 10 random treatment trajectories per patient and time step.

1242 F.1.1 VIOLATIONS OF THE POSITIVITY ASSUMPTION
1243

1244 We observed violations of the positivity (overlap) assumption in several instantiations of the
1245 semi-synthetic dataset generated with the parameters proposed by Melnychuk et al. (2022) and
1246 closely followed by several other works like (El Bouchattaoui et al., 2024; Wang et al., 2025).
1247 Concretely, for some random initializations almost all realized treatments are 0; for others, the dis-
1248 tribution is heavily skewed toward 1. Inspecting the individual (per-arm) propensities $p_{\ell,t}^A \in (0, 1)$
1249 defined by Eq. 38 reveals that a large fraction of values are effectively *degenerate*. For one seed, for
1250 example, 95.6% of per-arm propensities are $< 1\%$, 76.5% are $< 0.1\%$, 42.9% are $< 0.01\%$, 15.8%
1251 are $< 0.001\%$, and 2.9% are $< 0.0001\%$; only 28 out of 101,031 valid treatment decisions have
1252 propensity $> 50\%$. For another seed, the mass concentrates near 1: 8.7% of propensities exceed
1253 99% and 3.2% exceed 99.99%.

1254 While the positivity assumption requires $0 < \Pr(A_t = a \mid H_t) < 1$ almost surely, in practice causal
1255 estimators become unstable when a substantial mass of propensities lies outside $[\epsilon, 1 - \epsilon]$ for a small
1256 ϵ (e.g., 10^{-3}). The extreme values above arise because the *logit* in Eq. 38 (a linear combination of
1257 recent outcomes and covariate features) can be very large in magnitude for some seeds, pushing $\sigma(\cdot)$
1258 close to 0 or 1. In the next subsection we describe a minimally invasive modification that ensures
1259 overlap while preserving sequential confounding structure.

1260 F.2 OUR VERSION OF THE SEMI-SYNTHETIC DATASET
1261

1262 **Positivity via a monotone floor/ceiling.** To guarantee per-arm overlap we apply a monotone
1263 remapping to the *final* probability:

$$1264 \tilde{p}_{\ell,t}^A = q + (1 - 2q) \sigma(b_\ell + z_{\ell,t}), \quad q \in (0, 0.5), \quad (42)$$

1266 which forces $\tilde{p}_{\ell,t}^A \in [q, 1 - q]$. We use $q = 0.15$.

1268 **Preserving confounding via logit normalization.** A naive floor alone avoids practical violations
1269 of positivity assumption but can still yield weak dependence on confounders if the *logit* distribution
1270 collapses (e.g., is almost always very large or very small). We therefore re-scale the *pre-bias* logit
1271 using train-set statistics so that the sigmoid operates on a stable range:

$$1273 r_{\ell,t} = \gamma_\ell^Y \bar{Y}_{t-1} + \gamma_\ell^X f_X^{(\ell)}(X_t), \quad (43)$$

$$1274 z_{\ell,t} = \frac{r_{\ell,t} - \mu_\ell}{\sigma_\ell + \varepsilon}, \quad \varepsilon > 0, \quad (44)$$

1276 where (μ_ℓ, σ_ℓ) are the mean and standard deviation of $r_{\ell,t}$ estimated *on the training split only*. The
1277 final propensity is then given by Eq. 42. This is an affine, monotone transformation of the original
1278 logit and therefore preserves the ordering of $r_{\ell,t}$ with respect to the history H_t .

1280 **Two-pass generation to avoid leakage.** We use a standard two-pass protocol:
1281

1. **Pass 1 (train only, original policy).** We run the generator once using Eq. 38 and record
1282 $r_{\ell,t}$ from Eq. 43 for every (ℓ, t) on the training split. We compute (μ_ℓ, σ_ℓ) per arm via an
1283 online (Welford) estimator. The trajectories from this pass are discarded; only (μ_ℓ, σ_ℓ) are
1284 kept.
2. **Pass 2 (train/val/test, overlap-calibrated).** We regenerate all splits from scratch. At each
1285 step we recompute $r_{\ell,t}$ from the *current* pass’s history, apply the z-score in Eq. 44, then
1286 compute $\tilde{p}_{\ell,t}^A$ via Eq. 42 and sample treatments. Thus, sequential dependence on past out-
1287 comes/treatments remains intact; the first pass only provides (μ_ℓ, σ_ℓ) , analogous to feature
1288 normalization. The magnitude of the utilized bias term is sufficiently small to not make
1289 logit magnitudes too large.

1293 Pass 2 recomputes the logit from the *realized* past outcomes and treatments of the same pass; pass 1
1294 probabilities are never used for sampling. Since z-scoring is affine and the final mapping is mono-
1295 tone, the confounding signal (how H_t shifts treatment odds) is preserved, while the floor prevents
near-degenerate propensities that destabilize estimation and calibration.

1296 For one random instantiation of our new dataset, we have that the minimal probability of an in-
 1297 dividual treatment is 15.7%, and the maximum probability is 84.6%: apart from avoiding values
 1298 too close to 0% or to 100%, the sigmoid does not get completely saturated, which would produce
 1299 minimal or maximal values exactly in the floor. Apart from that, 86.7% of per-arm propensities are
 1300 > 25%, 14.3% are > 50%, and 1.1% are > 75%. For another seed, we have that the minimum
 1301 per-arm propensity score is 15.8% and the highest one is 84.9%. Also, we have 89.1% of per-arm
 1302 propensities > 25%, 40.7% > 50%, and 5.2% > 75%.

1303 1304 F.3 DETAILS ON EXPERIMENTS WITH REAL-WORLD DATA

1305 In line with the semi-synthetic setup (App. F.1), we rely on MIMIC-EXTRACT (Wang et al., 2020),
 1306 a standardized preprocessing pipeline for ICU time series (hourly resolution). Missing values are
 1307 imputed using forward and backward filling, and all continuous time-varying variables are stan-
 1308 dardized. We use the same set of $d_x = 25$ vital signs and the same three static attributes (gender,
 1309 ethnicity, age), one-hot encoded, yielding $d_v = 44$ static features. Both the time-varying covariates
 1310 and static features are treated as potential confounders.

1311 We consider $d_a = 2$ binary interventions: vasopressors and mechanical ventilation. The factual
 1312 outcome is diastolic blood pressure ($d_y = 1$). Clinically, both interventions can increase or decrease
 1313 blood pressure depending on context, motivating counterfactual trajectory analysis under alternative
 1314 treatment choices.

1315 **Cohort and splits.** We select 5,000 patients with ICU stays of at least 30 hours; stays are truncated
 1316 at 60 hours. The cohort is divided into train/validation/test sets with a 70%/15%/15% split.

1319 1320 G BASELINES

1321 1322 G.1 CAUSAL TRANSFORMER

1323 1324 G.1.1 BASE CAUSAL TRANSFORMER

1325 We implement the Causal Transformer (CT) of Melnychuk et al. (2022) as a strong baseline for
 1326 estimating

$$1327 \mathbb{E}[Y_{t+\tau}[\bar{a}_{t:t+\tau-1}] \mid \bar{H}_t] \quad (45)$$

1328 under a treatment plan $\bar{a}_{t:t+\tau-1}$. To avoid duplication, we reuse the multi-input transformer encoder
 1329 in App. B and highlight only CT-specific pieces (projection inputs, balanced-representation learning,
 1330 and stabilizers).

1331 **Inputs and autoregressive conditioning.** CT consumes three factual streams up to anchor time
 1332 t : covariates \bar{X}_t , outcomes \bar{Y}_t , and *left-shifted* treatments \bar{A}_{t-1} , plus static covariates V . For a
 1333 projection horizon τ , CT concatenates the factual histories with the (non-random) *future* intervention
 1334 sequence on the treatment stream and with *autoregressively fed predictions* on the outcome stream:

$$1335 \bar{A}_{t-1} \parallel \bar{a}_{t:t+\tau-1}, \quad (46)$$

$$1336 \bar{Y}_t \parallel \bar{Y}_{t+1:t+\tau-1}. \quad (47)$$

1337 Teacher forcing is used during training for multi-step prediction; at evaluation time, the model feeds
 1338 back its own predictions autoregressively. Static covariates V are injected in all subnetworks.

1339 **Architecture (encoder blocks, cross-attention, pooling).** CT follows the multi-input transformer
 1340 pattern in App. B: masked self-attention per stream, cross-attention between streams, position-wise
 1341 feed-forward layers, and LN+residual connections, with trainable relative positional encodings and
 1342 attentional dropout. After the last block, the three stream states are *averaged* and passed through a
 1343 Linear+ELU to obtain a balanced representation $\Phi_t \in \mathbb{R}^{d_r}$:

$$1344 \Phi_t = \text{ELU}(\text{Linear}(\frac{1}{3}(\mathbf{A}_{t-1}^B + \mathbf{X}_t^B + \mathbf{Y}_t^B))). \quad (48)$$

1345 (Implementation note: CT omits the final output projection after concatenating attention heads to
 1346 reduce overfitting.)

1350 **Balanced-representation training.** CT trains Φ_t to be (i) predictive of the one-step factual outcome while (ii) *non-predictive* of the current treatment with a Counterfactual Domain Confusion (CDC) loss. Two light heads are attached to Φ_t : an outcome head G_Y and a treatment classifier G_A . Let d_a be the number of treatment categories. The losses are

$$1355 \quad \mathcal{L}_{GA}(\theta_A, \theta_R) = - \sum_{j=1}^{d_a} \mathbf{1}\{A_t = a_j\} \log (G_A(\Phi_t(\theta_R); \theta_A)_j), \quad (49)$$

$$1359 \quad \mathcal{L}_{\text{conf}}(\theta_A, \theta_R) = - \sum_{j=1}^{d_a} \frac{1}{d_a} \log (G_A(\Phi_t(\theta_R); \theta_A)_j), \quad (50)$$

1362 and the alternating min–min scheme is

$$1364 \quad (\hat{\theta}_Y, \hat{\theta}_R) = \arg \min_{\theta_Y, \theta_R} \mathcal{L}_{GY}(\theta_Y, \theta_R) + \alpha \mathcal{L}_{\text{conf}}(\hat{\theta}_A, \theta_R), \quad (51)$$

$$1367 \quad \hat{\theta}_A = \arg \min_{\theta_A} \alpha \mathcal{L}_{GA}(\theta_A, \hat{\theta}_R), \quad (52)$$

1369 with $\alpha > 0$ the domain-confusion weight and \mathcal{L}_{GY} defined by the chosen outcome head (see below).

1371 **Training stabilizers and augmentation.** We follow CT practice: (i) an *exponential moving average* (EMA) of parameters across trainable modules; (ii) *attentional dropout*; and (iii) mini-batch 1372 augmentation that duplicates samples and randomly *masks* the last t_s covariate steps in the duplicate 1373 (to reflect unavailable future covariates for $\tau \geq 2$).

1376 **Point-estimator CT (original).** The original CT uses a point head G_Y with squared error:

$$1378 \quad \mathcal{L}_{GY}^{(\text{point})}(\theta_Y, \theta_R) = \|\mathbf{y}_{t+1} - G_Y(\Phi_t(\theta_R), \mathbf{a}_t; \theta_Y)\|_2^2. \quad (53)$$

1380 G.1.2 DISTRIBUTIONAL VARIANTS

1382 We additionally evaluate two *distributional* adaptations of CT that replace the outcome head/loss, 1383 keeping architecture and CDC unchanged.

1384 **CT–Gaussian head (heteroscedastic NLL).** The Gaussian head predicts per-dimension mean and 1385 variance $(\hat{\mu}_{t+1}, \hat{\sigma}_{t+1}^2) = G_Y^N(\Phi_t, \mathbf{a}_t)$, and minimizes the Gaussian negative log-likelihood (diagonal 1386 covariance):

$$1388 \quad \mathcal{L}_{GY}^N(\theta_Y, \theta_R) = \frac{1}{2} \left\| \frac{\mathbf{y}_{t+1} - \hat{\mu}_{t+1}}{\hat{\sigma}_{t+1}} \right\|_2^2 + \frac{1}{2} \mathbf{1}^\top \log \hat{\sigma}_{t+1}^2. \quad (54)$$

1392 **CT–CRPS / random-quantile head.** The random-quantile head predicts outcome quantiles given 1393 $\alpha \in (0, 1)^{d_y}$. Let $\hat{q}_{\alpha, j} = d_j(\Phi_t, \mathbf{a}_t, \alpha_j)$ denote the predicted α_j -quantile of $Y_{t+1, j}$ for branch j .

1394 Drawing K i.i.d. vectors $\{\alpha^{(k)}\}_{k=1}^K$ with entries $\alpha_j^{(k)} \sim \text{Unif}(0, 1)$, we use the Monte Carlo CRPS 1395 objective:

$$1397 \quad \mathcal{L}_{GY}^{\text{CRPS}}(\theta_Y, \theta_R) = \sum_{j=1}^{d_y} \frac{1}{K} \sum_{k=1}^K \rho_{\alpha_j^{(k)}}(y_{t+1, j} - \hat{q}_{\alpha_j^{(k)}, j}), \quad (55)$$

1400 with the pinball loss

$$1401 \quad \rho_\alpha(u) = (\alpha - \mathbf{1}\{u < 0\}) u. \quad (56)$$

1403 This is the same random-quantile reconstruction used for \mathbf{Y} in G-Latent, providing a proper scoring rule (CRPS) and capturing predictive uncertainty through the quantile function.

1404 G.2 G-NET
1405

1406 G-Net implements g-computation in two steps. First, it estimates the conditional *expectations* of
1407 within-time components of $\mathbf{L}_{t+1} = (\mathbf{Y}_{t+1}, \mathbf{X}_{t+1})$ given history and action. Concretely, for an
1408 ordered decomposition $\mathbf{L}_{t+1}^{(0)}, \dots, \mathbf{L}_{t+1}^{(p-1)}$, we learn

$$1409 \mathbb{E}[\mathbf{L}_{t+1}^{(j)} \mid \bar{\mathbf{H}}_t, \mathbf{A}_t, \mathbf{L}_{t+1}^{(0:j-1)}] \quad (57)$$

1410 with a two-layer LSTM. Samples from the corresponding conditionals are obtained by adding residuals
1411 drawn from an empirical error distribution built on a 10% holdout split (residual bootstrap).
1412 Training uses teacher forcing and an MSE loss.
1413

1414 Second, counterfactual trajectories under a treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$ are generated by Monte Carlo,
1415 rolling the learned conditionals forward across steps.

1416 We follow the same architecture class reported alongside CT: one–two layered LSTMs, a linear
1417 representation layer, and a small feed-forward head on top. At evaluation, we simulate under $\bar{\mathbf{a}}$ with
1418 start-of-interval indexing (action \mathbf{A}_s precedes $(\mathbf{Y}_{s+1}, \mathbf{X}_{s+1})$), using the residual-bootstrap sampler.
1419

1420 G.3 TRANSFORMER G-NET
1421

1422 Transformer G-Net follows the same two-step pipeline but replaces the recurrent modules with the
1423 multi-input transformer encoder of App. B. The transformer encodes the factual history before action
1424 into a fused state \mathbf{r}_t (respecting start-of-interval indexing). For an ordered within-time decomposi-
1425 tion $\mathbf{L}_{t+1}^{(0)}, \dots, \mathbf{L}_{t+1}^{(p-1)}$, each conditional expectation is predicted by a small MLP head conditioned
1426 on \mathbf{r}_t , \mathbf{A}_t , and previously generated groups; training uses teacher forcing and an MSE objective.
1427 During rollout we inject residual noise via the same 10% holdout bootstrap and obtain the *distribu-
1428 tion* at horizon $t+\tau$ as the empirical measure over M Monte Carlo trajectories (again $M=50$),
1429 without any balanced-representation objective.
1430

1431 H METRICS
1432

1433 Our model outputs i.i.d. Monte Carlo (MC) samples $\{\mathbf{y}_{t+s}^{(m,i)}\}_{m=1}^M$ from the interventional law
1434 $p_{\bar{\mathbf{a}}}(\mathbf{y}_{t+s} \mid \bar{\mathbf{h}}_t^{(i)})$ at each relative step $s \in \{1, \dots, \tau\}$, given history $\bar{\mathbf{h}}_t^{(i)}$ and a treatment plan $\bar{\mathbf{a}}_{t:t+\tau-1}$.
1435 All metrics are computed *per step* and averaged over n test patients; when relevant we also report a
1436 trajectory-level score aggregating all steps.
1437

1438 **RMSE of the predictive mean (point accuracy).** Let the per-step predictive mean for patient i
1439 be

$$1440 \hat{\mu}_{t+s}^{(i)} = \frac{1}{M} \sum_{m=1}^M \mathbf{y}_{t+s}^{(m,i)}. \quad (58)$$

1441 For a d_y -dimensional outcome we report

$$1442 \text{RMSE}_s = \sqrt{\frac{1}{n d_y} \sum_{i=1}^n \|\hat{\mu}_{t+s}^{(i)} - \mathbf{y}_{t+s}^{(i)}\|_2^2}, \quad (59)$$

1443 which summarizes point accuracy of the posterior mean implied by the predictive distribution (lower
1444 is better).
1445

1446 **KDE log-likelihood (density fit).** We estimate the patient-specific predictive density at relative
1447 step t' with an isotropic Gaussian KDE using a single global bandwidth $h > 0$:

$$1448 \hat{f}_{t'}^{(i)}(\mathbf{y}) = \frac{1}{M} \sum_{m=1}^M \mathcal{N}(\mathbf{y}; \mathbf{y}_{t+t'}^{(m,i)}, h^2 \mathbf{I}_{d_y}), \quad (60)$$

1449 where \mathbf{I}_{d_y} is the $d_y \times d_y$ identity matrix and h is fixed across all t' and all patients. The metric is
1450 the average log-likelihood:
1451

$$1452 \text{KDE-LL}_{t'} = \frac{1}{n} \sum_{i=1}^n \log \hat{f}_{t'}^{(i)}(\mathbf{y}_{t+t'}^{(i)}). \quad (61)$$

1458 **Energy score (multivariate proper scoring rule).** For $d_y \geq 1$, the energy score (ES) for a predictive distribution F_s and realization \mathbf{y}_{t+s} is
 1459
 1460

$$\text{ES}_s(F_s, \mathbf{y}_{t+s}) = \mathbb{E}[\|\mathbf{X} - \mathbf{y}_{t+s}\|_2] - \frac{1}{2} \mathbb{E}[\|\mathbf{X} - \mathbf{X}'\|_2], \quad (62)$$

1461 with $\mathbf{X}, \mathbf{X}' \sim F_s$ i.i.d. Using MC samples, we estimate
 1462
 1463

$$\widehat{\text{ES}}_s = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{1}{M} \sum_{m=1}^M \|\mathbf{y}_{t+s}^{(m,i)} - \mathbf{y}_{t+s}^{(i)}\|_2 - \frac{1}{2M(M-1)} \sum_{m \neq m'} \|\mathbf{y}_{t+s}^{(m,i)} - \mathbf{y}_{t+s}^{(m',i)}\|_2 \right\}, \quad (63)$$

1464 which is strictly proper and *sensitive to cross-dimensional dependence* (lower is better). In the
 1465 univariate case ($d_y=1$) ES equals the continuous ranked probability score (CRPS).
 1466

1467 **Global (pathwise) energy score (temporal coherence).** To assess coherence across *all* output
 1468 dimensions and steps, we compute ES on the concatenated outcome vector $\tilde{\mathbf{y}}_{t+1:t+\tau} \in \mathbb{R}^{\tau d_y}$, where
 1469 $\tilde{\mathbf{y}}_{t+1:t+\tau}^{(m,i)} := [\mathbf{y}_{t+1}^{(m,i)}; \dots; \mathbf{y}_{t+\tau}^{(m,i)}]$ and $\tilde{\mathbf{y}}_{t+1:t+\tau}^{(i)} := [\mathbf{y}_{t+1}^{(i)}; \dots; \mathbf{y}_{t+\tau}^{(i)}]$:
 1470

$$\text{GES} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{1}{M} \sum_{m=1}^M \|\tilde{\mathbf{y}}_{t+1:t+\tau}^{(m,i)} - \tilde{\mathbf{y}}_{t+1:t+\tau}^{(i)}\|_2 - \frac{1}{2M(M-1)} \sum_{m \neq m'} \|\tilde{\mathbf{y}}_{t+1:t+\tau}^{(m,i)} - \tilde{\mathbf{y}}_{t+1:t+\tau}^{(m',i)}\|_2 \right\}. \quad (64)$$

1471 This whole-trajectory ES rewards correct temporal correlations and cross-step consistency of the
 1472 joint predictive law (lower is better).
 1473

1474 **Quantile coverage (calibration).** For quantile levels $\mathcal{Q} = \{0.1, 0.2, \dots, 0.9\}$, we compare each
 1475 realized outcome component to the MC-estimated predictive quantile of that component. Let $(\cdot)_j$
 1476 denote the j -th component. Define
 1477

$$\hat{Q}_{s,j}^{(i)}(q) := \text{quantile}_q(\{\mathbf{y}_{t+s}^{(m,i)}\}_{m=1}^M). \quad (65)$$

1478 Per step and per dimension, the empirical q -coverage is
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$$\widehat{\text{Cov}}_{s,j}(q) = \frac{1}{n} \sum_{i=1}^n \mathbb{E} \left\{ (\mathbf{y}_{t+s}^{(i)})_j \leq \hat{Q}_{s,j}^{(i)}(q) \right\}, \quad (66)$$

1480 which should match the nominal level q for a calibrated model (higher/lower than q indicates over-/under-coverage). We use “ \leq ” to break ties; quantiles are computed from MC samples per (i, s, j)
 1481 with a fixed interpolation rule.
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1483 We also use aggregations across steps:
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$$\widehat{\text{Cov}}_j^{\text{steps}}(q) = \frac{1}{n \tau} \sum_{s=1}^{\tau} \sum_{i=1}^n \mathbb{E} \left\{ (\mathbf{y}_{t+s}^{(i)})_j \leq \hat{Q}_{s,j}^{(i)}(q) \right\}, \quad (67)$$

1485 **Calibration summary (MAE).** A scalar summary is the mean absolute calibration error, averaged
 1486 over quantiles, dimensions, and steps:
 1487

$$\text{CalMAE} = \frac{1}{|\mathcal{Q}| d_y \tau} \sum_{q \in \mathcal{Q}} \sum_{j=1}^{d_y} \sum_{s=1}^{\tau} \left| \widehat{\text{Cov}}_{s,j}(q) - q \right|. \quad (68)$$

1488 Lower is better; per-dimension or per-step variants follow by omitting the corresponding averages.
 1489

I HYPERPARAMETERS

I.1 MULTI-INPUT TRANSFORMER

1490 For better comparability, we used the same multi-input transformer hyperparameters for all the models
 1491 that use transformer processing (CT, CT-CRPS, Transformer G-Net, CT-Gaussian and G-Latent).
 1492 We used the same hyperparameters as Melnychuk et al. (2022), as additional tuning on our specific
 1493 models did not provide significant improvements. We list these hyperparameters in table 4, and
 1494 define them next:
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Table 4: Architectural hyperparameters for the multi-input transformer.

Hyperparameter	Semi-synthetic	Real-world
Transformer units	24	24
Representation size	44	22
Fully connected hidden units	22	22
Dropout rate	0.1	0.2
Transformer blocks	1	2
Attention heads	2	3
Max relative position	20	30

- Transformer units: model width per stream (token and attention projection size; per-head dimension roughly Transformer units divided by Attention heads).
- Representation size: fused history embedding dimension used downstream.
- Fully connected hidden units: inner width of the position-wise feed-forward sublayer.
- Dropout rate: probability used after linear layers in attention and feed-forward sublayers.
- Transformer blocks: number of stacked encoder blocks.
- Attention heads: number of heads in multi-head attention.
- Max relative position: clipping radius for relative positional encodings shared across blocks and streams.

I.2 CAUSAL TRANSFORMER

We report the specific training hyperparameters of CT in table 5.

Table 5: Training hyperparameters for the multi-input transformer.

Hyperparameter	Semi-synthetic	Real-world
Learning rate	0.01	0.0001
Batch size	64	64
Max epochs	400	300

For the distributional versions of CT, we used the same hyperparameters. For CT-CRPS, we used a number of α quantile MC samples $K = 5$ for both semi-synthetic and real-world dataset. This value is the same we used for G-Latent.

I.3 G-NET

For G-Net, we used the hyperparameters configuration from the implementation in Melnychuk et al. (2022). We report it in table 6, and define them as:

- Recurrent layers: number of stacked recurrent layers.
- Sequence hidden units: hidden size per recurrent layer.
- Fully connected hidden units: width of the feed-forward head.
- Dropout rate: dropout probability in recurrent/feed-forward parts.
- Representation size: size of the intermediate representation.
- Learning rate: optimizer step size.
- Batch size: examples per minibatch.
- Max epochs: maximum training epochs.

I.4 G-LATENT

As previously mentioned, the multi-input transformer we used in G-Latent has the hyperparameters shared with other baselines and defined in I.1. As for the rest of hyperparameters, after an optimization process based on factual validation datasets, we selected the ones shown in table 7. We defined next:

1566 Table 6: Architectural and training hyperparameters for G-Net.
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1568 Hyperparameter	1569 Semi-synthetic	1570 Real-world
1569 Recurrent layers	1	2
1570 Sequence hidden units	148	144
1571 Fully connected hidden units	74	72
1572 Dropout rate	0.1	0.1
1573 Representation size	74	72
1574 Learning rate	0.01	0.001
1575 Batch size	256	256
1575 Max epochs	200	200

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- 1578 Learning rate: optimizer step size.
- 1579 KL weight: coefficient on the KL divergence term in the ELBO.
- 1580 Latent dimension: dimensionality of the VAE latent variable z .
- 1581 Auxiliar loss weight (λ_{aux}): weight on the auxiliary one-step prediction loss.
- 1582 Max epochs: maximum number of training epochs.
- 1583 Reconstruction weights (outcome, covariates): multipliers for outcome and covariate reconstruction terms. The fact that, in both datasets, covariates have much higher coefficients than outcomes makes the model give balanced weight to both of them. Weights are selected in such a way that the sum of products of each weight with each dimensionality gives one.
- 1584 MC α samples (K): number of quantile levels sampled per step for the CRPS/quantile head.
- 1585 Batch size: number of examples per minibatch.
- 1586 Context dimension: size of the context vector fed to the VAE.
- 1587 Encoder hidden sizes: layer widths of the encoder MLP $q_\phi(z | x, c)$.
- 1588 Decoder hidden sizes: layer widths of the shared decoder trunk T_θ .
- 1589 Quantile-branch hidden sizes: layer widths in the per-outcome, α -aware branches.
- 1590 Shared decoder layers: count of initial decoder layers shared by the α -aware and mean/log-variance branches.
- 1591 Warm-up epochs (auxiliar loss only): epochs optimizing only the auxiliary loss before enabling VAE terms.
- 1592 GRU hidden size: hidden width of the temporal GRU cell used in latent rollouts.

1602 Table 7: Architectural and training hyperparameters for the RNN+Conditional VAE (G-Latent)
1603 stack.
1604

1605 Hyperparameter	1606 Semi-synthetic	1607 Real-world
1606 Learning rate	0.0001	0.0003
1607 KL weight	1.0	1.0
1608 Latent dimension	6	6
1609 Auxiliar loss weight (λ_{aux})	0.1	0.1
1610 Max epochs	70	110
1611 Reconstruction weights (outcome, covariates)	[6.67, 0.32]	[18.0, 0.32]
1612 MC α samples (K)	5	5
1613 Batch size	8	8
1614 Context dimension	256	256
1615 Encoder hidden sizes	[256, 256, 256, 256, 256]	[256, 256, 256, 256, 256]
1616 Decoder hidden sizes	[256, 256, 256, 256, 256]	[256, 256, 256, 256, 256]
1617 Quantile-branch hidden sizes	[64, 64]	[128, 128]
1618 Shared decoder layers	3	3
1619 Warm-up epochs (auxiliar loss only)	20	30
GRU hidden size	64	64

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J ADDITIONAL RESULTS

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J.1 SEMI-SYNTHETIC DATASET (OUR NEW VERSION)

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In table 8, we show the Energy Scores for our new modified semi-synthetic dataset. In tables 9, 10 and 11 we show the KDE-LL for bandwidths 0.2, 0.3 and 0.4, respectively. In table 12, we show the RMSE metrics. Finally, in tables 13 and 14 we show the empirical quantile coverage for all the steps (1 to 6 in the first table, 7 to 11 in the second one, plus across step coverage), the dimensions, and several quantiles from 0.1 to 0.9. The bolded results are the ones closest to the expected coverage percentage, i.e., for quantile 0.1, 10%, for quantile 0.2, 20%, etc.

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Table 8: Energy Score per step t' on semi-synthetic dataset (corrected benchmark). Rightmost column reports the Global Energy Score across steps. Best per column in **bold**.

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Model	$t'=1$	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$	Global
G-Net	0.17 ± 0.00	0.39 ± 0.03	0.39 ± 0.04	0.45 ± 0.04	0.51 ± 0.05	0.55 ± 0.06	0.59 ± 0.06	0.63 ± 0.07	0.65 ± 0.07	0.68 ± 0.07	0.70 ± 0.08	1.85 ± 0.20
Transformer G-Net	0.37 ± 0.04	0.34 ± 0.04	0.40 ± 0.05	0.46 ± 0.06	0.50 ± 0.07	0.53 ± 0.08	0.56 ± 0.10	0.58 ± 0.11	0.60 ± 0.12	0.62 ± 0.13	0.64 ± 0.14	1.71 ± 0.11
CT-CRPS	0.09 ± 0.01	0.26 ± 0.06	0.32 ± 0.07	0.37 ± 0.07	0.41 ± 0.07	0.45 ± 0.08	0.48 ± 0.07	0.50 ± 0.07	0.53 ± 0.07	0.55 ± 0.07	0.57 ± 0.07	1.52 ± 0.23
CT-Gaussian	0.09 ± 0.01	0.25 ± 0.06	0.30 ± 0.07	0.34 ± 0.08	0.37 ± 0.08	0.40 ± 0.09	0.42 ± 0.09	0.44 ± 0.09	0.46 ± 0.09	0.48 ± 0.09	0.49 ± 0.09	1.35 ± 0.29
D.S. G-VAE (Gaussian)	0.28 ± 0.01	0.40 ± 0.02	0.49 ± 0.04	0.54 ± 0.05	0.58 ± 0.06	0.60 ± 0.06	0.62 ± 0.07	0.64 ± 0.07	0.65 ± 0.07	0.66 ± 0.07	0.67 ± 0.07	2.01 ± 0.20
D.S. G-VAE (CRPS)	0.13 ± 0.00	0.23 ± 0.04	0.28 ± 0.05	0.32 ± 0.06	0.35 ± 0.06	0.38 ± 0.06	0.40 ± 0.07	0.42 ± 0.07	0.44 ± 0.06	0.45 ± 0.06	0.47 ± 0.06	1.28 ± 0.21
G-Latent (Gaussian)	0.31 ± 0.02	0.35 ± 0.03	0.38 ± 0.04	0.40 ± 0.05	0.42 ± 0.05	0.44 ± 0.06	0.45 ± 0.06	0.46 ± 0.06	0.47 ± 0.06	0.48 ± 0.06	0.48 ± 0.06	1.51 ± 0.18
G-Latent (CRPS)	0.19 ± 0.02	0.25 ± 0.04	0.29 ± 0.05	0.33 ± 0.06	0.35 ± 0.06	0.37 ± 0.07	0.39 ± 0.07	0.40 ± 0.07	0.42 ± 0.07	0.42 ± 0.08	0.43 ± 0.08	1.25 ± 0.23

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Table 9: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.2. Best per column in **bold**.

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Model	$t'=1$	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	0.30 ± 0.05	-0.85 ± 0.20	-1.48 ± 0.25	-1.91 ± 0.28	-2.21 ± 0.31	-2.47 ± 0.33	-2.70 ± 0.36	-2.89 ± 0.39	-3.04 ± 0.42	-3.17 ± 0.45	-3.29 ± 0.48
Transformer G-Net	-1.34 ± 0.20	-1.07 ± 0.24	-1.52 ± 0.27	-1.86 ± 0.38	-2.12 ± 0.49	-2.33 ± 0.59	-2.52 ± 0.69	-2.70 ± 0.79	-2.86 ± 0.90	-3.00 ± 0.98	-3.14 ± 1.06
CT-CRPS	0.99 ± 0.07	-0.88 ± 0.71	-1.56 ± 0.78	-2.16 ± 0.82	-2.67 ± 0.79	-3.17 ± 0.81	-3.63 ± 0.81	-4.06 ± 0.79	-4.43 ± 0.81	-4.78 ± 0.83	-5.08 ± 0.86
CT-Gaussian	1.00 ± 0.05	-0.40 ± 0.49	-0.73 ± 0.54	-0.99 ± 0.57	-1.20 ± 0.57	-1.39 ± 0.57	-1.56 ± 0.57	-1.72 ± 0.57	-1.86 ± 0.57	-2.01 ± 0.57	-2.14 ± 0.57
D.S. G-VAE (Gaussian)	-1.21 ± 0.08	-1.87 ± 0.10	-2.25 ± 0.14	-2.45 ± 0.17	-2.57 ± 0.18	-2.65 ± 0.19	-2.70 ± 0.20	-2.74 ± 0.20	-2.76 ± 0.20	-2.79 ± 0.20	-2.80 ± 0.20
D.S. G-VAE (CRPS)	0.45 ± 0.07	-0.32 ± 0.33	-0.66 ± 0.37	-0.89 ± 0.39	-1.06 ± 0.40	-1.22 ± 0.40	-1.34 ± 0.39	-1.45 ± 0.38	-1.54 ± 0.36	-1.62 ± 0.33	-1.69 ± 0.31
G-Latent (Gaussian)	-1.36 ± 0.11	-1.52 ± 0.16	-1.62 ± 0.18	-1.69 ± 0.20	-1.74 ± 0.21	-1.79 ± 0.22	-1.83 ± 0.23	-1.86 ± 0.23	-1.88 ± 0.23	-1.90 ± 0.23	-1.92 ± 0.23
G-Latent (CRPS)	-0.01 ± 0.17	-0.50 ± 0.30	-0.78 ± 0.35	-0.98 ± 0.39	-1.12 ± 0.41	-1.24 ± 0.42	-1.34 ± 0.44	-1.42 ± 0.44	-1.48 ± 0.44	-1.53 ± 0.43	-1.59 ± 0.44

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Table 10: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.3. Best per column in **bold**.

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Model	$t'=1$	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	-0.02 ± 0.02	-0.79 ± 0.15	-1.26 ± 0.20	-1.59 ± 0.23	-1.84 ± 0.24	-2.04 ± 0.26	-2.22 ± 0.27	-2.36 ± 0.29	-2.48 ± 0.30	-2.58 ± 0.31	-2.67 ± 0.33
Transformer G-Net	-1.09 ± 0.19	-1.00 ± 0.22	-1.34 ± 0.23	-1.58 ± 0.30	-1.77 ± 0.36	-1.93 ± 0.42	-2.06 ± 0.40	-2.18 ± 0.54	-2.29 ± 0.60	-2.38 ± 0.65	-2.47 ± 0.69
CT-CRPS	0.38 ± 0.04	-0.63 ± 0.40	-1.02 ± 0.46	-1.35 ± 0.49	-1.64 ± 0.48	-1.90 ± 0.49	-2.15 ± 0.48	-2.38 ± 0.48	-2.57 ± 0.48	-2.76 ± 0.49	-2.92 ± 0.49
CT-Gaussian	0.37 ± 0.03	-0.51 ± 0.38	-0.75 ± 0.40	-0.94 ± 0.43	-1.09 ± 0.44	-1.23 ± 0.45	-1.34 ± 0.46	-1.45 ± 0.46	-1.55 ± 0.46	-1.64 ± 0.46	-1.73 ± 0.46
D.S. G-VAE (Gaussian)	-1.29 ± 0.07	-1.88 ± 0.08	-2.23 ± 0.12	-2.49 ± 0.15	-2.54 ± 0.17	-2.62 ± 0.18	-2.67 ± 0.18	-2.71 ± 0.19	-2.73 ± 0.19	-2.75 ± 0.19	-2.77 ± 0.19
D.S. G-VAE (CRPS)	0.08 ± 0.04	-0.47 ± 0.27	-0.72 ± 0.31	-0.90 ± 0.33	-1.04 ± 0.34	-1.16 ± 0.34	-1.27 ± 0.34	-1.35 ± 0.33	-1.43 ± 0.32	-1.50 ± 0.30	-1.56 ± 0.29
G-Latent (Gaussian)	-1.41 ± 0.09	-1.54 ± 0.13	-1.63 ± 0.15	-1.69 ± 0.17	-1.74 ± 0.18	-1.79 ± 0.19	-1.82 ± 0.20	-1.85 ± 0.20	-1.87 ± 0.20	-1.89 ± 0.20	-1.91 ± 0.20
G-Latent (CRPS)	-0.24 ± 0.12	-0.59 ± 0.22	-0.80 ± 0.27	-0.96 ± 0.30	-1.06 ± 0.32	-1.18 ± 0.33	-1.26 ± 0.34	-1.32 ± 0.34	-1.38 ± 0.34	-1.42 ± 0.34	-1.47 ± 0.34

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1681Table 11: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.4. Best per column in **bold**.

Model	$t'=1$	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	-0.37 ± 0.01	-0.91 ± 0.12	-1.27 ± 0.17	-1.53 ± 0.19	-1.74 ± 0.21	-1.91 ± 0.22	-2.06 ± 0.24	-2.18 ± 0.25	-2.28 ± 0.26	-2.37 ± 0.27	-2.45 ± 0.28
Transformer G-Net	-0.12 ± 0.17	-1.19 ± 0.19	-1.35 ± 0.21	-1.54 ± 0.26	-1.69 ± 0.31	-1.83 ± 0.35	-1.92 ± 0.41	-2.01 ± 0.45	-2.10 ± 0.49	-2.17 ± 0.53	-2.24 ± 0.56
CT-CRPS	-0.12 ± 0.02	-0.75 ± 0.26	-1.00 ± 0.30	-1.22 ± 0.33	-1.40 ± 0.34	-1.57 ± 0.35	-1.73 ± 0.35	-1.87 ± 0.35	-2.00 ± 0.35	-2.11 ± 0.36	-2.22 ± 0.36
CT-Gaussian	-0.13 ± 0.02	-0.73 ± 0.26	-0.91 ± 0.31	-1.05 ± 0.34	-1.17 ± 0.35	-1.27 ± 0.37	-1.36 ± 0.37	-1.44 ± 0.38	-1.51 ± 0.38	-1.58 ± 0.38	-1.64 ± 0.38
D.S. G-VAE (Gaussian)	-1.40 ± 0.06	-1.98 ± 0.07	-2.26 ± 0.12	-2.45 ± 0.14	-2.56 ± 0.16	-2.63 ± 0.17	-2.68 ± 0.17	-2.72 ± 0.18	-2.74 ± 0.18	-2.76 ± 0.18	-2.78 ± 0.18
D.S. G-VAE (CRPS)	-0.30 ± 0.03	-0.70 ± 0.21	-0.89 ± 0.25	-1.03 ± 0.27	-1.14 ± 0.29	-1.24 ± 0.29	-1.33 ± 0.29	-1.40 ± 0.29	-1.46 ± 0.28	-1.52 ± 0.27	-1.58 ± 0.26
G-Latent (Gaussian)	-1.51 ± 0.08	-1.62 ± 0.11	-1.70 ± 0.14	-1.76 ± 0.15	-1.80 ± 0.16	-1.84 ± 0.17	-1.87 ± 0.18	-1.90 ± 0.18	-1.92 ± 0.18	-1.94 ± 0.18	-1.95 ± 0.18
G-Latent (CRPS)	-0.54 ± 0.09	-0.78 ± 0.17	-0.95 ± 0.21	-1.08 ± 0.24	-1.18 ± 0.26	-1.26 ± 0.27	-1.32 ± 0.28	-1.37 ± 0.29	-1.42 ± 0.29	-1.46 ± 0.29	-1.50 ± 0.29

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1683Table 12: RMSE per step t' on semi-synthetic dataset (corrected benchmark). Best per column in **bold**.

Model	$t'=1$	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	0.28 ± 0.01	0.51 ± 0.05	0.64 ± 0.07	0.74 ± 0.08	0.81 ± 0.09	0.88 ± 0.09	0.94 ± 0.10	0.98 ± 0.11	1.02 ± 0.11	1.06 ± 0.12	1.09 ± 0.12
Transformer G-Net	0.60 ± 0.06	0.56 ± 0.06	0.66 ± 0.08	0.74 ± 0.10	0.80 ± 0.13	0.84 ± 0.15	0.89 ± 0.17	0.92 ± 0.19	0.95 ± 0.21	0.98 ± 0.22	1.00 ± 0.23
CT-Gaussian	0.17 ± 0.02	0.46 ± 0.11	0.54 ± 0.13	0.60 ± 0.14	0.64 ± 0.14	0.68 ± 0.14	0.71 ± 0.14	0.74 ± 0.14	0.76 ± 0.14	0.79 ± 0.14	0.81 ± 0.14
CT-CRPS	0.16 ± 0.02	0.48 ± 0.10	0.58 ± 0.11	0.65 ± 0.11	0.71 ± 0.10	0.76 ± 0.10	0.80 ± 0.10	0.84 ± 0.10	0.87 ± 0.10	0.89 ± 0.10	0.92 ± 0.10
CT	0.14 ± 0.01	0.34 ± 0.07	0.43 ± 0.10	0.49 ± 0.11	0.53 ± 0.12	0.56 ± 0.13	0.58 ± 0.13	0.60 ± 0.13	0.62 ± 0.13	0.64 ± 0.13	0.65 ± 0.13
D.S. G-VAE (Gaussian)	0.26 ± 0.00	0.44 ± 0.06	0.54 ± 0.09	0.61 ± 0.10	0.66 ± 0.11	0.70 ± 0.12	0.73 ± 0.12	0.76 ± 0.13	0.79 ± 0.13	0.81 ± 0.13	0.83 ± 0.13
D.S. G-VAE (CRPS)	0.23 ± 0.00	0.40 ± 0.07	0.49 ± 0.10	0.55 ± 0.11	0.59 ± 0.12	0.63 ± 0.12	0.66 ± 0.12	0.69 ± 0.12	0.72 ± 0.12	0.74 ± 0.12	0.76 ± 0.12
G-Latent (Gaussian)	0.35 ± 0.03	0.46 ± 0.06	0.53 ± 0.09	0.58 ± 0.10	0.61 ± 0.11	0.64 ± 0.12	0.67 ± 0.12	0.69 ± 0.12	0.71 ± 0.12	0.72 ± 0.12	0.73 ± 0.12
G-Latent (CRPS)	0.33 ± 0.04	0.44 ± 0.07	0.51 ± 0.10	0.60 ± 0.12	0.63 ± 0.12	0.66 ± 0.13	0.68 ± 0.13	0.70 ± 0.13	0.71 ± 0.13	0.73 ± 0.13	0.73 ± 0.13

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1692Table 13: Empirical coverage (%) by step and dimension for each quantile q . Steps $t' \in \{1, \dots, 6\}$, two outcome dimensions.

Model	Step t'											
	1		2		3		4		5		6	
	Dim 1	Dim 2										
Quantile $q = 0.1$												
G-Net	11.54 ± 2.61	12.61 ± 3.20	14.04 ± 3.53	15.62 ± 4.43	15.33 ± 4.63	17.32 ± 4.86	16.19 ± 5.38	18.38 ± 5.14	16.86 ± 5.73	19.17 ± 5.29	17.33 ± 5.97	19.82 ± 5.32
Transformer G-Net	23.52 ± 9.91	24.92 ± 5.68	11.51 ± 6.44	16.22 ± 4.97	11.87 ± 7.14	18.83 ± 4.85	12.50 ± 7.51	20.09 ± 4.54	13.21 ± 7.90	22.26 ± 4.31	13.88 ± 8.18	23.55 ± 4.20
CT-CRPS	8.78 ± 2.77	15.43 ± 6.74	16.47 ± 1.73	21.24 ± 7.52	17.94 ± 2.02	25.14 ± 8.93	19.28 ± 2.47	28.42 ± 9.75	20.30 ± 2.60	30.79 ± 10.73	21.08 ± 2.77	32.92 ± 11.52
CT-Gaussian	13.91 ± 6.81	15.34 ± 5.09	9.59 ± 2.68	13.11 ± 2.40	11.08 ± 3.35	16.69 ± 3.56	12.44 ± 4.19	19.45 ± 4.16	13.45 ± 4.95	21.59 ± 4.36	14.09 ± 5.50	23.30 ± 4.57
D.S. G-VAE (Gaussian)	0.19 ± 0.03	0.07 ± 0.05	0.72 ± 0.38	0.19 ± 0.19	0.71 ± 0.40	0.16 ± 0.17	0.67 ± 0.38	0.14 ± 0.15	0.66 ± 0.40	0.13 ± 0.14	0.68 ± 0.43	0.14 ± 0.14
D.S. G-VAE (CRPS)	0.20 ± 0.03	0.22 ± 0.02	0.55 ± 0.03	0.21 ± 0.04	0.77 ± 0.27	0.10 ± 0.07	0.80 ± 0.24	0.10 ± 0.07	10.00 ± 2.20	0.13 ± 0.09	0.00 ± 0.07	11.79 ± 8.41
G-Latent (Gaussian)	0.47 ± 0.07	0.11 ± 0.10	1.38 ± 0.51	0.41 ± 0.42	1.49 ± 0.52	0.30 ± 0.61	2.44 ± 1.02	0.82 ± 0.74	2.70 ± 1.13	0.00 ± 0.88	1.11 ± 1.23	1.13 ± 0.95
G-Latent (CRPS)	8.89 ± 1.94	9.48 ± 3.75	0.88 ± 1.77	10.23 ± 2.93	10.37 ± 1.88	10.86 ± 2.97	10.79 ± 1.94	11.46 ± 3.35	11.17 ± 1.95	11.96 ± 3.70	11.45 ± 2.04	12.64 ± 3.88
Quantile $q = 0.2$												
G-Net	20.30 ± 4.64	30.87 ± 5.44	31.03 ± 4.54	32.74 ± 5.34	31.83 ± 5.70	38.78 ± 5.56	34.24 ± 5.30	34.53 ± 6.03	32.79 ± 5.87	31.35 ± 4.33	30.66 ± 8.83	32.09 ± 4.41
Transformer G-Net	22.46 ± 8.68	42.59 ± 4.71	29.37 ± 6.17	33.49 ± 5.98	28.68 ± 7.22	35.88 ± 5.44	37.62 ± 4.32	29.54 ± 8.34	39.11 ± 4.39	30.19 ± 8.69	40.30 ± 4.69	37.46 ± 13.70
CT-CRPS	15.22 ± 4.00	23.94 ± 9.75	33.21 ± 1.82	29.47 ± 10.16	24.37 ± 2.21	32.93 ± 11.42	25.58 ± 2.66	36.14 ± 12.25	26.62 ± 2.80	38.33 ± 13.10	27.46 ± 3.06	40.18 ± 13.70
CT-Gaussian	23.29 ± 9.12	25.21 ± 5.91	16.84 ± 2.59	44.26 ± 5.60	18.65 ± 5.30	26.19 ± 4.67	28.47 ± 6.47	31.20 ± 5.71	21.30 ± 4.62	30.93 ± 4.85	22.14 ± 7.56	32.63 ± 4.84
D.S. G-VAE (Gaussian)	0.76 ± 0.15	0.34 ± 0.18	2.21 ± 0.75	0.70 ± 0.61	2.33 ± 0.83	0.69 ± 0.62	2.51 ± 0.88	0.73 ± 0.63	2.65 ± 0.92	0.75 ± 0.63	2.81 ± 0.99	0.80 ± 0.66
D.S. G-VAE (CRPS)	0.85 ± 0.16	0.25 ± 0.26	16.85 ± 6.39	14.61 ± 3.60	20.78 ± 6.02	14.01 ± 3.60	44.83 ± 2.44	30.47 ± 3.55	34.29 ± 4.16	19.83 ± 4.57	25.22 ± 4.80	20.38 ± 5.22
G-Latent (Gaussian)	1.27 ± 1.45	2.35 ± 1.24	1.72 ± 1.26	30.28 ± 3.34	30.76 ± 2.98	20.77 ± 2.06	36.01 ± 3.24	31.45 ± 3.27	20.92 ± 2.76	34.32 ± 3.14	31.23 ± 3.14	31.23 ± 3.24
G-Latent (CRPS)	7.17 ± 1.35	2.17 ± 1.18	1.74 ± 1.19	2.36 ± 1.49	2.17 ± 1.49	2.34 ± 1.49	2.15 ± 1.50	2.37 ± 1.50	2.16 ± 1.50	2.37 ± 1.50	2.27 ± 1.50	2.27 ± 1.50
Quantile $q = 0.3$												
G-Net	39.36 ± 4.56	39.96 ± 5.35	31.03 ± 4.54	32.74 ± 5.34	31.83 ± 5.70	38.78 ± 5.56	34.24 ± 5.30	34.53 ± 6.03	32.79 ± 5.87	31.35 ± 4.33	30.11 ± 8.69	35.39 ± 6.13
Transformer G-Net	44.92 ± 4.73	42.55 ± 4.71	42.45 ± 4.71	48.31 ± 4.37	39.37 ± 6.17	38.73 ± 5.44	37.62 ± 4.32	29.54 ± 8.34	39.11 ± 4.39	30.19 ± 8.69	40.30 ± 4.69	37.46 ± 13.70
CT-CRPS	20.46 ± 4.77	34.04 ± 4.02	30.13 ± 4.32	30.13 ± 4.32	30.13 ± 4.32	32.93 ± 4.28	31.64 ± 4.28	42.56 ± 4.77	34.25 ± 2.77	44.45 ± 4.77	33.22 ± 3.40	46.00 ± 14.33
CT-Gaussian	31.99 ± 4.51	34.65 ± 4.02	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	32.93 ± 4.28	40.46 ± 5.02
D.S. G-VAE (Gaussian)	3.76 ± 1.38	2.64 ± 0.99	1.51 ± 0.53	3.47 ± 1.64	7.43 ± 1.49	3.63 ± 1.75	7.98 ± 1.62	3.94 ± 1.81	8.53 ± 1.71	4.24 ± 1.84	8.07 ± 1.81	4.53 ± 2.08
D.S. G-VAE (CRPS)	22.96 ± 4.69	18.36 ± 2.97	27.09 ± 7.03	24.62 ± 4.17	30.46 ± 6.15	26.69 ± 4.69	32.77 ± 5.57	28.19 ± 5.54	34.29 ± 5.12	29.09 ± 4.60	35.35 ± 5.03	29.55 ± 7.24
G-Latent (Gaussian)	7.37 ± 1.35	4.11 ± 1.18	1.54 ± 1.21	5.24 ± 3.43	13.44 ± 1.94	4.81 ± 1.59	12.43 ± 2.16	5.06 ± 1.59	16.13 ± 2.24	4.70 ± 1.59	12.23 ± 2.34	4.81 ± 2.24
G-Latent (CRPS)	28.14 ± 3.58	26.52 ± 6.23	28.71 ± 3.42	28.72 ± 4.76	28.97 ± 3.39	29.51 ± 4.49	29.09 ± 3.49	30.30 ± 4.61	29.24 ± 3.73	40.76 ± 5.17	39.29 ± 3.90	41.26 ± 5.14
Quantile $q = 0.4$												
G-Net	40.05 ± 4.66	49.09 ± 5.64	47.81 ± 3.51	49.36 ± 4.69	47.71 ± 3.99	49.05 ± 5.73	47.59 ± 4.32	49.41 ± 5.24	49.08 ± 4.70	40.08 ± 5.16	40.25 ± 5.45	42.47 ± 5.87
Transformer G-Net	56.04 ± 6.10	53.33 ± 3.34	58.31 ± 5.24	42.16 ± 4.20	37.68 ± 7.25	45.46 ± 5.07	46.42 ± 4.71	51.99 ± 4.27	46.48 ± 7.08	52.83 ± 4.91	46.67 ± 7.58	55.09 ± 4.48
CT-CRPS	5.47 ± 1.37	56.28 ± 18										

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1737 Table 14: Empirical coverage (%) by step and dimension for each quantile q . Steps $t' \in \{7, \dots, 11\}$
1738 and calibration across steps, two outcome dimensions.

Model	Step t' & Across Steps (Cal.)												
	7		8		9		10		11		Across Steps (Cal.)		
	Dim 1	Dim 2											
Quantile $q = 0.1$													
G-Net	17.62 \pm 6.07	20.20 \pm 5.19	17.90 \pm 6.30	20.53 \pm 5.11	17.95 \pm 6.32	20.51 \pm 4.96	18.03 \pm 6.28	20.46 \pm 4.85	18.11 \pm 6.38	20.44 \pm 4.75	16.94 \pm 5.62	19.24 \pm 4.93	
Transformer-G-Net	14.53 \pm 8.45	24.71 \pm 4.14	15.12 \pm 8.75	25.69 \pm 4.24	15.76 \pm 8.98	26.30 \pm 4.23	16.26 \pm 9.25	26.83 \pm 4.25	16.65 \pm 9.55	27.32 \pm 4.26	14.13 \pm 8.14	23.24 \pm 4.10	
CT-G-VAE	21.48 \pm 6.45	34.42 \pm 4.22	20.51 \pm 6.35	35.22 \pm 4.22	20.73 \pm 6.35	36.73 \pm 4.25	20.26 \pm 6.35	37.06 \pm 4.28	20.51 \pm 6.35	37.34 \pm 4.28	19.30 \pm 6.30	22.86 \pm 4.34	
CT-Gaussian	14.86 \pm 6.16	24.59 \pm 4.66	15.21 \pm 6.63	25.93 \pm 4.95	16.19 \pm 7.25	27.04 \pm 5.11	16.05 \pm 7.80	28.03 \pm 5.40	17.54 \pm 8.18	28.81 \pm 5.66	14.17 \pm 6.60	22.86 \pm 4.32	
D.S. G-VAE (Gaussian)	0.71 \pm 0.43	0.13 \pm 0.14	0.75 \pm 0.43	0.13 \pm 0.14	0.79 \pm 0.48	0.13 \pm 0.14	0.84 \pm 0.49	0.12 \pm 0.15	0.90 \pm 0.53	0.14 \pm 0.15	0.74 \pm 0.42	0.14 \pm 0.15	
D.S. G-VAE (CRPS)	15.37 \pm 3.71	11.89 \pm 3.53	15.79 \pm 3.81	12.08 \pm 3.88	16.01 \pm 3.95	11.97 \pm 4.07	16.24 \pm 4.27	11.80 \pm 4.19	16.51 \pm 4.56	11.65 \pm 4.32	14.10 \pm 3.79	10.77 \pm 2.98	
G-Latent (Gaussian)	3.21 \pm 1.30	1.28 \pm 1.04	3.41 \pm 1.40	1.40 \pm 1.05	3.57 \pm 1.45	1.55 \pm 1.14	3.71 \pm 1.54	1.64 \pm 1.11	3.85 \pm 1.66	1.77 \pm 1.11	2.92 \pm 1.19	1.16 \pm 0.90	
G-Latent (CRPS)	11.69 \pm 2.11	13.19 \pm 4.05	11.89 \pm 2.20	13.60 \pm 4.14	12.09 \pm 2.36	13.80 \pm 4.20	12.23 \pm 2.45	14.00 \pm 4.14	12.40 \pm 2.57	14.15 \pm 4.13	11.40 \pm 2.05	12.59 \pm 3.72	
Quantile $q = 0.2$													
G-Net	25.81 \pm 6.25	28.45 \pm 5.88	26.06 \pm 6.41	28.67 \pm 5.70	26.19 \pm 6.64	28.61 \pm 5.60	26.27 \pm 6.68	28.51 \pm 5.54	26.36 \pm 6.84	28.51 \pm 5.43	25.25 \pm 5.85	27.54 \pm 5.60	
Transformer-G-Net	22.77 \pm 9.13	33.83 \pm 4.60	23.41 \pm 9.34	34.83 \pm 4.82	24.17 \pm 9.57	35.35 \pm 4.94	24.62 \pm 9.82	35.84 \pm 5.04	25.13 \pm 10.11	35.32 \pm 5.19	22.44 \pm 8.65	32.29 \pm 4.36	
CT-G-VAE	29.46 \pm 6.44	41.28 \pm 4.22	28.45 \pm 6.42	42.05 \pm 4.22	29.35 \pm 6.46	42.36 \pm 4.24	29.02 \pm 6.42	42.78 \pm 4.26	29.57 \pm 6.42	43.07 \pm 4.26	27.47 \pm 6.39	32.80 \pm 4.32	
CT-Gaussian	23.00 \pm 8.24	33.78 \pm 4.86	23.67 \pm 8.72	34.97 \pm 5.01	24.54 \pm 9.24	35.01 \pm 5.29	25.25 \pm 9.86	36.89 \pm 5.68	25.96 \pm 10.26	37.07 \pm 5.92	22.18 \pm 7.63	32.04 \pm 4.75	
D.S. G-VAE (Gaussian)	2.99 \pm 0.99	0.87 \pm 0.70	3.17 \pm 1.03	0.93 \pm 0.74	3.35 \pm 1.05	1.02 \pm 0.82	3.54 \pm 1.08	1.06 \pm 0.82	3.69 \pm 1.08	1.11 \pm 0.86	2.92 \pm 0.74	0.87 \pm 0.70	
D.S. G-VAE (CRPS)	26.01 \pm 4.90	20.77 \pm 5.83	26.49 \pm 4.98	21.01 \pm 6.42	26.78 \pm 5.39	20.77 \pm 6.69	27.01 \pm 5.84	20.54 \pm 6.93	27.19 \pm 6.27	20.25 \pm 7.08	24.22 \pm 4.91	19.39 \pm 4.99	
G-Latent (Gaussian)	8.57 \pm 2.05	4.77 \pm 2.38	9.05 \pm 2.14	5.35 \pm 2.46	9.36 \pm 2.24	5.74 \pm 2.57	9.60 \pm 2.30	6.11 \pm 2.61	9.87 \pm 2.45	6.48 \pm 2.67	7.85 \pm 1.89	4.35 \pm 2.16	
G-Latent (CRPS)	20.49 \pm 3.34	22.31 \pm 4.32	20.73 \pm 3.58	22.83 \pm 4.40	20.82 \pm 3.57	22.83 \pm 4.46	20.94 \pm 3.72	22.92 \pm 4.46	21.06 \pm 3.77	22.94 \pm 4.45	20.12 \pm 3.12	21.51 \pm 4.12	
Quantile $q = 0.3$													
G-Net	33.22 \pm 5.95	35.69 \pm 6.01	33.39 \pm 6.13	35.89 \pm 5.84	33.55 \pm 6.41	35.77 \pm 5.72	33.65 \pm 6.54	35.62 \pm 5.65	33.71 \pm 6.71	35.57 \pm 5.61	32.87 \pm 5.62	35.00 \pm 5.75	
Transformer-G-Net	30.85 \pm 8.99	41.31 \pm 5.60	31.51 \pm 9.18	41.29 \pm 5.48	32.15 \pm 9.45	42.67 \pm 5.70	32.63 \pm 9.67	41.30 \pm 5.92	33.11 \pm 9.91	41.59 \pm 6.16	30.70 \pm 8.37	39.96 \pm 4.63	
CT-G-VAE	35.49 \pm 6.35	47.88 \pm 4.22	35.87 \pm 6.45	48.05 \pm 4.22	36.34 \pm 6.56	48.78 \pm 4.24	36.34 \pm 6.56	49.34 \pm 4.24	36.57 \pm 6.62	49.41 \pm 4.24	35.49 \pm 6.35	49.55 \pm 4.24	
CT-Gaussian	31.16 \pm 8.28	41.48 \pm 5.58	31.71 \pm 10.04	42.87 \pm 5.93	32.55 \pm 10.46	43.46 \pm 5.20	32.37 \pm 10.97	44.50 \pm 5.74	32.85 \pm 11.31	45.15 \pm 6.08	30.27 \pm 8.99	40.38 \pm 5.54	
D.S. G-VAE (Gaussian)	9.38 \pm 1.90	4.85 \pm 2.16	9.88 \pm 1.93	5.20 \pm 2.29	19.51 \pm 3.85	24.54 \pm 4.03	5.48 \pm 2.52	10.18 \pm 2.06	5.67 \pm 2.62	10.98 \pm 2.13	4.93 \pm 2.84	4.69 \pm 2.15	
D.S. G-VAE (CRPS)	36.16 \pm 6.06	29.84 \pm 8.01	36.69 \pm 5.99	29.99 \pm 8.68	36.93 \pm 6.38	29.51 \pm 9.97	37.13 \pm 6.87	29.05 \pm 9.36	37.24 \pm 7.42	26.62 \pm 9.57	34.41 \pm 5.36	28.52 \pm 7.00	
G-Latent (Gaussian)	17.66 \pm 2.38	13.29 \pm 3.36	18.25 \pm 2.44	14.22 \pm 3.49	18.60 \pm 2.52	14.78 \pm 3.56	18.88 \pm 2.55	19.59 \pm 3.74	18.31 \pm 2.32	19.21 \pm 2.71	15.78 \pm 3.59	16.55 \pm 2.27	
G-Latent (CRPS)	29.63 \pm 4.28	31.95 \pm 4.66	29.86 \pm 4.41	32.49 \pm 4.56	29.91 \pm 4.52	32.29 \pm 4.63	29.96 \pm 4.54	32.22 \pm 4.72	29.93 \pm 4.56	32.07 \pm 4.76	29.48 \pm 3.85	31.18 \pm 4.56	
Quantile $q = 0.4$													
G-Net	40.31 \pm 5.57	42.66 \pm 5.80	40.44 \pm 5.76	42.81 \pm 5.64	40.53 \pm 6.01	42.64 \pm 5.51	40.62 \pm 6.23	42.44 \pm 5.37	40.76 \pm 6.48	42.33 \pm 5.37	40.20 \pm 5.24	42.29 \pm 5.53	
Transformer-G-Net	38.96 \pm 9.46	47.98 \pm 5.60	39.53 \pm 9.18	48.78 \pm 5.48	41.10 \pm 9.01	49.19 \pm 6.49	40.47 \pm 9.30	49.56 \pm 6.57	40.85 \pm 9.57	50.00 \pm 7.10	39.02 \pm 7.87	40.95 \pm 5.03	
CT-G-VAE	39.50 \pm 6.33	49.43 \pm 6.99	40.09 \pm 10.54	50.21 \pm 6.72	40.49 \pm 10.83	51.23 \pm 6.56	41.33 \pm 7.01	51.53 \pm 7.01	41.99 \pm 11.64	51.99 \pm 8.10	38.85 \pm 6.63	48.22 \pm 6.73	
CT-Gaussian	23.60 \pm 6.66	36.44 \pm 3.78	45.79 \pm 7.70	46.14 \pm 3.99	45.83 \pm 7.97	45.38 \pm 2.88	45.82 \pm 8.20	44.84 \pm 4.50	45.81 \pm 8.29	43.27 \pm 5.51	46.49 \pm 4.66	46.49 \pm 3.06	
D.S. G-VAE (Gaussian)	45.81 \pm 7.60	46.44 \pm 3.78	45.79 \pm 7.70	46.14 \pm 3.99	45.83 \pm 7.97	45.48 \pm 2.88	45.82 \pm 8.20	44.84 \pm 4.50	45.81 \pm 8.29	43.22 \pm 5.51	46.49 \pm 4.66	46.49 \pm 3.06	
D.S. G-VAE (CRPS)	55.53 \pm 5.55	47.81 \pm 10.99	56.80 \pm 5.80	47.62 \pm 11.81	56.04 \pm 6.44	46.74 \pm 12.39	56.02 \pm 7.05	47.12 \pm 12.83	56.49 \pm 7.54	47.13 \pm 13.17	54.77 \pm 4.43	47.51 \pm 10.16	
G-Latent (Gaussian)	47.55 \pm 1.36	48.42 \pm 3.72	47.48 \pm 1.46	48.76 \pm 3.76	47.33 \pm 1.61	47.18 \pm 1.82	48.11 \pm 3.58	46.91 \pm 1.88	47.18 \pm 3.61	48.06 \pm 4.63	50.93 \pm 5.01	49.54 \pm 3.69	
G-Latent (CRPS)	49.23 \pm 2.40	52.00 \pm 4.94	49.18 \pm 4.43	52.21 \pm 4.84	49.01 \pm 4.53	51.81 \pm 4.81	48.66 \pm 4.63	51.29 \pm 4.86	48.48 \pm 4.81	50.93 \pm 5.01	49.54 \pm 3.69	51.62 \pm 5.12	
Quantile $q = 0.5$													
G-Net	54.67 \pm 5.57	56.87 \pm 4.88	54.54 \pm 4.81	56.89 \pm 4.18	54.50 \pm 5.10	56.58 \pm 4.72	54.36 \pm 5.87	56.55 \pm 4.87	56.33 \pm 5.94	54.48 \pm 5.61	56.16 \pm 3.86	54.94 \pm 4.25	56.89 \pm 4.02
Transformer-G-Net	54.57 \pm 7.41	60.53 \pm 4.41	55.87 \pm 7.77	61.08 \pm 7.12	56.11 \pm 8.22	61.34 \pm 7.92	56.32 \pm 9.67	61.27 \pm 7.97	57.39 \pm 9.09	61.55 \pm 7.97	56.63 \pm 8.30	60.22 \pm 5.86	60.22 \pm 5.86
CT-G-VAE	57.53 \pm 6.33	59.47 \pm 6.80	57.76 \pm 9.48	63.81 \pm 10.17	58.10 \pm 9.63	60.47 \pm 10.36	58.85 \pm 9.93	63.48 \pm 10.30	59.01 \pm 10.83	57.63 \pm 8.53	59.22 \pm 9.26	58.85 \pm 8.06	58.85 \pm 8.06
CT-Gaussian													

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J.2 SEMI-SYNTHETIC DATASET (ORIGINAL VERSION)

1784 In table 15, we show a summary of results for selected steps for the original semi-synthetic dataset with issues
1785 regarding the positivity assumption. In table 16, we show the Energy Scores. In tables 17, 18 and 19 we show
1786 the KDE-LL for bandwidths 0.2, 0.3 and 0.4, respectively. In table 20, we show the RMSE metrics.1787 Table 15: Results at selected steps $t' \in \{3, 5, 8, 11\}$ for the semi-synthetic dataset. Metrics: Energy
1788 Score (ES \downarrow) (per step and across steps), KDE-Loglikelihood (KDE-LL \uparrow), and RMSE \downarrow .
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Model	$t' = 3$			$t' = 5$			$t' = 8$			$t' = 11$			Global
	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	
G-Net	0.65 \pm 0.08	-2.22 \pm 0.39	0.82 \pm 0.04	0.99 \pm 0.11	-3.54 \pm 0.43	1.02 \pm 0.05	1.27 \pm 0.11	-4.65 \pm 0.44	1.22 \pm 0.06	1.41 \pm 0.14	-5.15 \pm 0.47	1.35 \pm 0.06	3.57 \pm 0.43
Transformer G-Net	0.49 \pm 0.08	-1.49 \pm 0.32	0.66 \pm 0.04	0.74 \pm 0.11	-2.42 \pm 0.43	0.80 \pm 0.04	1.10 \pm 0.11	-3.69 \pm 0.39	1.00 \pm 0.06	1.31 \pm 0.14	-4.14 \pm 0.35	1.17 \pm 0.06	2.92 \pm 0.38
CT (CRPS)	0.41 \pm 0.06	-1.40 \pm 0.20	0.67 \pm 0.06	0.53 \pm 0.06	-1.86 \pm 0.25	0.80 \pm 0.05	0.65 \pm 0.06	-2.29 \pm 0.24	0.94 \pm 0.06	0.73 \pm 0.05	-2.60 \pm 0.22	1.05 \pm 0.06	1.85 \pm 0.22
CT (Gaussian)	0.52 \pm 0.07	-1.56 \pm 0.32	0.64 \pm 0.06	0.65 \pm 0.06	-1.81 \pm 0.30	0.78 \pm 0.05	0.82 \pm 0.07	-2.19 \pm 0.29	0.91 \pm 0.05	0.93 \pm 0.07	-2.52 \pm 0.28	1.03 \pm 0.06	2.40 \pm 0.28
CT	0.46 \pm 0.01	0.51 \pm 0.02	0.55 \pm 0.02	0.61 \pm 0.02	...
D.S. G-VAE (Gaussian)	0.46 \pm 0.05	-2.39 \pm 0.40	0.89 \pm 0.05	1.00 \pm 0.05	-2.80 \pm 0.40	0.86 \pm 0.06	1.00 \pm 0.07	-3.01 \pm 0.34	0.84 \pm 0.07	0.92 \pm 0.08	-3.02 \pm 0.34	1.25 \pm 0.07	3.44 \pm 0.41
D.S. G-VAE (CRPS)	0.44 \pm 0.06	-1.57 \pm 0.25	0.68 \pm 0.04	0.51 \pm 0.05	-1.80 \pm 0.22	0.85 \pm 0.06	0.58 \pm 0.07	-2.01 \pm 0.24	1.10 \pm 0.07	0.65 \pm 0.08	-2.24 \pm 0.21	1.20 \pm 0.09	1.88 \pm 0.21
G-Latent (Gaussian)	0.40 \pm 0.04	-1.48 \pm 0.31	0.62 \pm 0.05	0.46 \pm 0.04	-1.66 \pm 0.26	0.70 \pm 0.05	0.51 \pm 0.05	-1.81 \pm 0.24	0.78 \pm 0.05	0.54 \pm 0.05	-1.91 \pm 0.24	0.83 \pm 0.07	1.64 \pm 0.13
G-Latent (CRPS)	0.39 \pm 0.06	-1.32 \pm 0.15	0.65 \pm 0.06	0.46 \pm 0.06	-1.59 \pm 0.16	0.77 \pm 0.06	0.53 \pm 0.06	-1.82 \pm 0.15	0.88 \pm 0.04	0.56 \pm 0.05	-1.95 \pm 0.14	0.94 \pm 0.03	1.67 \pm 0.20

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1796 Table 16: Energy Score per step t' on semi-synthetic dataset. Rightmost column reports the Global
1797 Energy Score across steps. Best per column in **bold**.

Model	$t' = 2$	$t' = 3$	$t' = 4$	$t' = 5$	$t' = 6$	$t' = 7$	$t' = 8$	$t' = 9$	$t' = 10$	$t' = 11$	Global
	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	
G-Net	0.44 \pm 0.06	0.65 \pm 0.08	0.84 \pm 0.09	0.99 \pm 0.11	1.11 \pm 0.11	1.20 \pm 0.13	1.27 \pm 0.11	1.33 \pm 0.09	1.38 \pm 0.13	1.41 \pm 0.14	3.57 \pm 0.43
Transformer G-Net	0.39 \pm 0.06	0.49 \pm 0.08	0.62 \pm 0.06	0.74 \pm 0.11	1.03 \pm 0.13	1.10 \pm 0.11	1.19 \pm 0.13	1.25 \pm 0.13	1.31 \pm 0.14	2.92 \pm 0.38	...
CT-CRPS	0.35 \pm 0.05	0.41 \pm 0.06	0.49 \pm 0.06	0.53 \pm 0.06	0.58 \pm 0.06	0.62 \pm 0.06	0.65 \pm 0.06	0.68 \pm 0.06	0.71 \pm 0.06	0.73 \pm 0.05	1.85 \pm 0.22
CT-Gaussian	0.42 \pm 0.05	0.52 \pm 0.07	0.58 \pm 0.07	0.65 \pm 0.06	0.72 \pm 0.06	0.75 \pm 0.07	0.82 \pm 0.07	0.88 \pm 0.08	0.91 \pm 0.07	0.93 \pm 0.07	2.40 \pm 0.28
D.S. G-VAE (Gaussian)	0.40 \pm 0.05	0.49 \pm 0.05	0.54 \pm 0.04	0.60 \pm 0.05	0.67 \pm 0.06	0.70 \pm 0.06	0.72 \pm 0.06	0.74 \pm 0.05	0.76 \pm 0.07	0.78 \pm 0.07	2.21 \pm 0.24
D.S. G-VAE (CRPS)	0.38 \pm 0.05	0.44 \pm 0.06	0.48 \pm 0.06	0.51 \pm 0.05	0.54 \pm 0.06	0.56 \pm 0.06	0.58 \pm 0.07	0.61 \pm 0.06	0.63 \pm 0.07	0.65 \pm 0.08	1.85 \pm 0.21
G-Latent (Gaussian)	0.36 \pm 0.04	0.40 \pm 0.04	0.43 \pm 0.04	0.46 \pm 0.04	0.48 \pm 0.04	0.49 \pm 0.05	0.51 \pm 0.05	0.52 \pm 0.05	0.53 \pm 0.05	0.54 \pm 0.05	1.64 \pm 0.13
G-Latent (CRPS)	0.34 \pm 0.05	0.39 \pm 0.06	0.43 \pm 0.06	0.46 \pm 0.06	0.49 \pm 0.06	0.51 \pm 0.06	0.53 \pm 0.06	0.54 \pm 0.06	0.55 \pm 0.06	0.56 \pm 0.05	1.67 \pm 0.20

1803

1804 Table 17: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.2. Best per
1805 column in **bold**.
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Model	$t' = 2$	$t' = 3$	$t' = 4$	$t' = 5$	$t' = 6$	$t' = 7$	$t' = 8$	$t' = 9$	$t' = 10$	$t' = 11$	Global
	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	KDE-LL \uparrow	RMSE \downarrow	ES \downarrow	
G-Net	-1.68 \pm 0.38	-3.09 \pm 0.45	-4.43 \pm 0.41	-5.71 \pm 0.61	-6.83 \pm 0.56	-7.72 \pm 0.40	-8.41 \pm 0.73	-8.99 \pm 0.60	-9.38 \pm 0.69	-9.69 \pm 0.71	...
Transformer G-Net	-1.24 \pm 0.31	-2.01 \pm 0.39	-2.68 \pm 0.32	-3.31 \pm 0.43	-3.80 \pm 0.50	-4.79 \pm 0.73	-5.63 \pm 0.56	-6.41 \pm 0.68	-7.12 \pm 0.81	-7.88 \pm 0.70	...
CT-CRPS	-1.13 \pm 0.20	-1.44 \pm 0.23	-1.69 \pm 0.32	-1.86 \pm 0.30	-2.01 \pm 0.37	-2.21 \pm 0.35	-2.38 \pm 0.39	-2.53 \pm 0.45	-2.67 \pm 0.31	-2.80 \pm 0.29	...
CT-Gaussian	-1.29 \pm 0.22	-1.67 \pm 0.27	-1.81 \pm 0.33	-1.95 \pm 0.37	-2.10 \pm 0.29	-2.29 \pm 0.33	-2.52 \pm 0.25	-2.72 \pm 0.30	-2.89 \pm 0.36	-3.02 \pm 0.41	...
D.S. G-VAE (Gaussian)	-1.90 \pm 0.34	-2.27 \pm 0.31	-2.51 \pm 0.45	-2.67 \pm 0.39	-2.79 \pm 0.31	-2.87 \pm 0.44	-2.94 \pm 0.41	-2.99 \pm 0.48	-3.04 \pm 0.51	-3.07 \pm 0.43	...
D.S. G-VAE (CRPS)	-1.26 \pm 0.25	-1.51 \pm 0.32	-1.70 \pm 0.21	-1.82 \pm 0.27	-1.95 \pm 0.31	-2.06 \pm 0.30	-2.16 \pm 0.39	-2.25 \pm 0.35	-2.33 \pm 0.42	-2.40 \pm 0.29	...
G-Latent (Gaussian)	-1.30 \pm 0.29	-1.53 \pm 0.18	-1.71 \pm 0.13	-1.83 \pm 0.11	-1.93 \pm 0.14	-2.02 \pm 0.19	-2.09 \pm 0.25	-2.15 \pm 0.30	-2.20 \pm 0.36	-2.26 \pm 0.41	...
G-Latent (CRPS)	-0.97 \pm 0.24	-1.31 \pm 0.27	-1.55 \pm 0.29	-1.72 \pm 0.30	-1.87 \pm 0.31	-1.98 \pm 0.31	-2.07 \pm 0.31	-2.14 \pm 0.30	-2.19 \pm 0.30	-2.25 \pm 0.29	...

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Table 18: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.3. Best per column in **bold**.

Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	-1.41 ± 0.32	-2.47 ± 0.39	-3.37 ± 0.34	-4.15 ± 0.46	-4.80 ± 0.49	-5.29 ± 0.54	-5.68 ± 0.40	-5.99 ± 0.46	-6.21 ± 0.41	-6.37 ± 0.51
Transformer G-Net	-1.23 ± 0.30	-1.83 ± 0.39	-2.37 ± 0.39	-3.12 ± 0.41	-3.81 ± 0.37	-4.43 ± 0.41	-4.44 ± 0.39	-4.68 ± 0.38	-5.25 ± 0.32	-5.77 ± 0.34
CT-CRPS	-1.02 ± 0.22	-1.29 ± 0.18	-1.52 ± 0.24	-1.71 ± 0.25	-1.88 ± 0.24	-2.01 ± 0.21	-2.14 ± 0.22	-2.22 ± 0.24	-2.26 ± 0.23	-2.30 ± 0.23
CT-Gaussian	-1.29 ± 0.31	-1.49 ± 0.38	-1.77 ± 0.32	-1.98 ± 0.39	-2.17 ± 0.30	-2.30 ± 0.37	-2.46 ± 0.34	-2.50 ± 0.29	-2.72 ± 0.28	-2.85 ± 0.31
D.S. G-VAE (Gaussian)	-1.92 ± 0.29	-2.27 ± 0.31	-2.49 ± 0.34	-2.65 ± 0.31	-2.76 ± 0.37	-2.85 ± 0.29	-2.91 ± 0.32	-2.96 ± 0.34	-3.00 ± 0.35	-3.03 ± 0.38
D.S. G-VAE (CRPS)	-1.29 ± 0.20	-1.50 ± 0.25	-1.66 ± 0.23	-1.77 ± 0.29	-1.87 ± 0.22	-1.97 ± 0.24	-2.05 ± 0.25	-2.12 ± 0.31	-2.19 ± 0.27	-2.26 ± 0.27
G-Latent (Gaussian)	-1.25 ± 0.37	-1.43 ± 0.30	-1.56 ± 0.25	-1.65 ± 0.22	-1.73 ± 0.21	-1.80 ± 0.20	-1.85 ± 0.20	-1.89 ± 0.20	-1.93 ± 0.21	-1.97 ± 0.23
G-Latent (CRPS)	-0.98 ± 0.18	-1.25 ± 0.20	-1.44 ± 0.21	-1.58 ± 0.21	-1.69 ± 0.21	-1.78 ± 0.21	-1.85 ± 0.20	-1.91 ± 0.20	-1.96 ± 0.19	-2.00 ± 0.19

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Table 19: KDE Loglikelihood per step t' on semi-synthetic dataset with bandwidth 0.4. Best per column in **bold**.

Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	-1.37 ± 0.32	-2.22 ± 0.39	-2.94 ± 0.39	-3.54 ± 0.43	-4.01 ± 0.44	-4.38 ± 0.42	-4.65 ± 0.44	-4.88 ± 0.45	-5.04 ± 0.40	-5.15 ± 0.47
Transformer G-Net	-1.21 ± 0.30	-1.49 ± 0.32	-1.89 ± 0.40	-2.42 ± 0.43	-2.91 ± 0.44	-3.38 ± 0.45	-3.69 ± 0.39	-3.85 ± 0.37	-4.01 ± 0.40	-4.14 ± 0.35
CT-CRPS	-1.15 ± 0.19	-1.40 ± 0.20	-1.62 ± 0.24	-1.86 ± 0.25	-1.99 ± 0.26	-2.14 ± 0.24	-2.29 ± 0.24	-2.35 ± 0.22	-2.48 ± 0.23	-2.60 ± 0.22
CT-Gaussian	-1.41 ± 0.31	-1.56 ± 0.32	-1.68 ± 0.34	-1.81 ± 0.30	-1.95 ± 0.29	-2.06 ± 0.31	-2.19 ± 0.29	-2.33 ± 0.28	-2.38 ± 0.28	-2.52 ± 0.28
D.S. G-VAE (Gaussian)	-1.97 ± 0.28	-2.30 ± 0.30	-2.51 ± 0.28	-2.66 ± 0.32	-2.77 ± 0.31	-2.85 ± 0.31	-2.91 ± 0.35	-2.96 ± 0.33	-2.99 ± 0.32	-3.02 ± 0.32
D.S. G-VAE (CRPS)	-1.39 ± 0.22	-1.57 ± 0.25	-1.70 ± 0.22	-1.80 ± 0.22	-1.89 ± 0.22	-1.97 ± 0.25	-2.04 ± 0.24	-2.11 ± 0.22	-2.18 ± 0.21	-2.24 ± 0.21
G-Latent (Gaussian)	-1.34 ± 0.36	-1.48 ± 0.31	-1.59 ± 0.28	-1.66 ± 0.26	-1.72 ± 0.25	-1.77 ± 0.24	-1.81 ± 0.24	-1.85 ± 0.24	-1.88 ± 0.24	-1.91 ± 0.24
G-Latent (CRPS)	-1.10 ± 0.14	-1.32 ± 0.15	-1.48 ± 0.16	-1.59 ± 0.16	-1.69 ± 0.16	-1.76 ± 0.16	-1.82 ± 0.15	-1.87 ± 0.15	-1.91 ± 0.14	-1.95 ± 0.14

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Table 20: RMSE per step t' on semi-synthetic dataset. Best per column in **bold**.

Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	$t'=7$	$t'=8$	$t'=9$	$t'=10$	$t'=11$
G-Net	0.67 ± 0.03	0.82 ± 0.04	0.96 ± 0.04	1.02 ± 0.05	1.09 ± 0.05	1.18 ± 0.05	1.22 ± 0.06	1.25 ± 0.06	1.29 ± 0.06	1.35 ± 0.06
Transformer G-Net	0.59 ± 0.03	0.66 ± 0.04	0.73 ± 0.04	0.80 ± 0.04	0.86 ± 0.05	0.92 ± 0.05	1.00 ± 0.06	1.06 ± 0.06	1.11 ± 0.06	1.17 ± 0.06
CT-Gaussian	0.54 ± 0.05	0.64 ± 0.06	0.72 ± 0.06	0.78 ± 0.05	0.84 ± 0.05	0.88 ± 0.06	0.91 ± 0.05	0.95 ± 0.05	0.99 ± 0.06	1.03 ± 0.06
CT-CRPS	0.55 ± 0.05	0.67 ± 0.06	0.76 ± 0.06	0.80 ± 0.05	0.85 ± 0.05	0.91 ± 0.06	0.94 ± 0.06	0.97 ± 0.06	1.02 ± 0.05	1.05 ± 0.06
CT	0.37 ± 0.01	0.46 ± 0.01	0.49 ± 0.01	0.51 ± 0.02	0.53 ± 0.02	0.54 ± 0.02	0.55 ± 0.02	0.58 ± 0.02	0.60 ± 0.02	0.61 ± 0.02
D.S. G-VAE (Gaussian)	0.56 ± 0.06	0.69 ± 0.05	0.79 ± 0.06	0.88 ± 0.07	0.97 ± 0.07	1.09 ± 0.06	1.18 ± 0.08	1.24 ± 0.09	1.30 ± 0.08	1.35 ± 0.08
D.S. G-VAE (CRPS)	0.57 ± 0.05	0.68 ± 0.06	0.77 ± 0.06	0.85 ± 0.06	0.93 ± 0.08	1.02 ± 0.06	1.10 ± 0.07	1.16 ± 0.09	1.22 ± 0.08	1.26 ± 0.09
G-Latent (Gaussian)	0.54 ± 0.05	0.62 ± 0.05	0.67 ± 0.05	0.70 ± 0.05	0.73 ± 0.05	0.76 ± 0.05	0.78 ± 0.05	0.79 ± 0.06	0.81 ± 0.07	0.83 ± 0.07
G-Latent (CRPS)	0.56 ± 0.06	0.65 ± 0.06	0.72 ± 0.06	0.77 ± 0.06	0.81 ± 0.05	0.85 ± 0.05	0.88 ± 0.04	0.90 ± 0.04	0.92 ± 0.03	0.94 ± 0.03

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J.3 REAL WORLD DATASET

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1892 In table 21, we show the Energy Scores. In table 22, we show the KDE-LL metric for bandwidth 3.6. In table
1893 23, we show the RMSE metrics.

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1895 Table 21: Energy Score per step t' on real-world dataset. Rightmost column reports the Global
1896 Energy Score across steps. Best per column in **bold**.

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Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$	Global
G-Net	5.32 ± 0.08	5.82 ± 0.08	6.29 ± 0.08	6.98 ± 0.09	7.44 ± 0.11	18.35 ± 0.33
Transformer G-Net	5.28 ± 0.06	5.84 ± 0.08	6.17 ± 0.09	6.47 ± 0.08	6.90 ± 0.08	16.70 ± 0.23
CT-CRPS	4.92 ± 0.06	5.39 ± 0.08	5.60 ± 0.07	5.77 ± 0.08	5.86 ± 0.07	14.61 ± 0.27
CT-Gaussian	5.25 ± 0.06	5.71 ± 0.08	5.99 ± 0.08	6.15 ± 0.07	6.34 ± 0.08	15.55 ± 0.23
D.S. G-VAE (Gaussian)	5.51 ± 0.08	5.99 ± 0.08	6.21 ± 0.10	6.34 ± 0.06	6.44 ± 0.07	15.98 ± 0.23
D.S. G-VAE (CRPS)	4.89 ± 0.08	5.36 ± 0.08	5.56 ± 0.09	5.70 ± 0.07	5.82 ± 0.06	14.38 ± 0.19
G-Latent (Gaussian)	5.27 ± 0.06	5.64 ± 0.08	5.84 ± 0.09	5.96 ± 0.07	6.07 ± 0.07	15.21 ± 0.26
G-Latent (CRPS)	4.85 ± 0.05	5.25 ± 0.08	5.47 ± 0.06	5.60 ± 0.09	5.72 ± 0.06	14.23 ± 0.23

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1906 Table 22: KDE Loglikelihood per step t' on real-world dataset with bandwidth 3.6. Best per column
1907 in **bold**.

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Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$
G-Net	-3.92 ± 0.05	-4.11 ± 0.05	-4.29 ± 0.06	-4.55 ± 0.07	-4.83 ± 0.04
Transformer G-Net	-3.89 ± 0.06	-4.06 ± 0.08	-4.16 ± 0.06	-4.30 ± 0.06	-4.48 ± 0.04
CT-CRPS	-3.81 ± 0.06	-3.94 ± 0.06	-3.99 ± 0.07	-4.08 ± 0.04	-4.19 ± 0.06
CT-Gaussian	-3.92 ± 0.06	-4.04 ± 0.07	-4.09 ± 0.06	-4.18 ± 0.06	-4.24 ± 0.07
D.S. G-VAE (Gaussian)	-3.90 ± 0.06	-3.98 ± 0.06	-4.01 ± 0.05	-4.03 ± 0.05	-4.04 ± 0.05
D.S. G-VAE (CRPS)	-3.82 ± 0.06	-3.92 ± 0.05	-3.94 ± 0.05	-3.99 ± 0.06	-4.04 ± 0.06
G-Latent (Gaussian)	-3.85 ± 0.06	-3.89 ± 0.06	-3.92 ± 0.05	-3.94 ± 0.04	-3.95 ± 0.06
G-Latent (CRPS)	-3.79 ± 0.06	-3.88 ± 0.05	-3.91 ± 0.05	-3.94 ± 0.05	-3.96 ± 0.06

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Table 23: RMSE per step t' on real-world dataset. Best per column in **bold**.

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Model	$t'=2$	$t'=3$	$t'=4$	$t'=5$	$t'=6$
G-Net	11.84 ± 0.24	12.83 ± 0.29	13.54 ± 0.33	14.05 ± 0.30	14.23 ± 0.29
Transformer G-Net	10.90 ± 0.30	11.67 ± 0.26	12.39 ± 0.38	12.96 ± 0.32	13.21 ± 0.29
CT-CRPS	9.34 ± 0.25	10.10 ± 0.29	10.53 ± 0.26	10.75 ± 0.29	10.91 ± 0.28
CT-Gaussian	9.63 ± 0.25	10.41 ± 0.29	10.74 ± 0.29	11.01 ± 0.34	11.25 ± 0.30
CT	9.00 ± 0.23	9.57 ± 0.24	9.90 ± 0.25	10.16 ± 0.27	10.35 ± 0.31
D.S. G-VAE (Gaussian)	9.58 ± 0.25	10.29 ± 0.22	10.66 ± 0.29	10.88 ± 0.26	11.04 ± 0.29
D.S. G-VAE (CRPS)	9.40 ± 0.22	10.09 ± 0.25	10.41 ± 0.23	10.63 ± 0.29	10.79 ± 0.30
G-Latent (Gaussian)	9.42 ± 0.23	10.09 ± 0.23	10.43 ± 0.25	10.64 ± 0.19	10.80 ± 0.25
G-Latent (CRPS)	9.23 ± 0.20	9.79 ± 0.24	10.14 ± 0.23	10.36 ± 0.29	10.55 ± 0.28

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K LLMs USAGE

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We used LLMs for diverse tasks in the production of this work. Mainly, for text and math reviewing and
correction. To a lesser extent, for discussing ideas.

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