
REINFORCE Converges to Optimal Policies with Any Learning Rate

Samuel Robertson^{1*} Thang D. Chu^{1*} Bo Dai^{2,3} Dale Schuurmans^{1,2}

Csaba Szepesvari^{1,2} Jincheng Mei²

¹University of Alberta ²Google DeepMind ³Georgia Institute of Technology

{smrobert,thang}@ualberta.ca
{bodai,schuurmans,szepi,jcmei}@google.com

Abstract

We prove that the classic REINFORCE stochastic policy gradient (SPG) method converges to globally optimal policies in finite-horizon Markov Decision Processes (MDPs) with *any* constant learning rate. To avoid the need for small or decaying learning rates, we introduce two key innovations in the stochastic bandit setting, which we then extend to MDPs. **First**, we identify a new exploration property of SPG: the online SPG method samples every action infinitely often (i.o.), improving on previous results that only guaranteed at least two actions would be sampled i.o. This means SPG inherently achieves asymptotic exploration without modification. **Second**, we eliminate the assumption of unique mean reward values, a condition that previous convergence analyses in the bandit setting relied on, but that does not translate to MDPs. Our results deepen the theoretical understanding of SPG in both bandit problems and MDPs, with a focus on how it handles the exploration-exploitation trade-off when standard optimization and stochastic approximation methods cannot be applied, as is the case with large constant learning rates.

1 Introduction

Policy gradient (PG) methods constitute one of the most popular classes of algorithms for reinforcement learning (RL). In the PG paradigm, a learner acts according to a parameterized policy; the expected return is directly optimized by computing its gradient with respect to the policy parameters and performing stochastic gradient ascent. PG methods have played a key role in the advancements of deep RL (Lillicrap et al., 2019; Schulman et al., 2017a,b): combined with deep neural networks, PG algorithms have shown strong empirical performance across many domains, including robotics Akkaya et al. (2019), games Vinyals et al. (2019), and large language model training (Rafailov et al., 2024; Ouyang et al., 2022).

Despite PG methods’ conceptual simplicity and rich set of practical applications, known theoretical guarantees on their performance come with restrictive assumptions. In particular, convergence proofs either require oracle access to the exact gradient (Liu et al., 2024; Agarwal et al., 2020), which is akin to demanding that the reward function and dynamics of the environment are known to the learner, or they impose harsh constraints on the learning rate used for stochastic gradient ascent (Mei et al., 2024b; Klein et al., 2024). Both of these assumptions are violated in typical applications. In the stochastic setting, where the rewards and transition probabilities are unknown and must be estimated from interaction with the environment, convergence of the classic REINFORCE algorithm (Williams, 1992) has only been shown under the assumption that the learning rate is either sufficiently small

*Equal contributions

(Klein et al., 2024) or decaying (Zhang et al., 2020). In this work we study REINFORCE with the standard softmax parameterization, and narrow the gap between theory and practice by providing the first proof in the stochastic setting that REINFORCE will globally converge to an optimal policy in tabular finite-horizon Markov Decision Processes (MDPs) with *any* constant learning rate. Along the way we show new results about the stochastic gradient bandit algorithm (Sutton and Barto, 2018; Mei et al., 2024a), which is the special case induced by applying REINFORCE to a bandit problem. Specifically, we show that the stochastic gradient bandit algorithm automatically achieves sufficient exploration for global convergence with an arbitrary constant learning rate; in doing so we remove a key assumption in prior work, and thus resolve an open problem posed by Mei et al. (2024a). Our results in the bandit setting extend to a more general “nonstationary bandit problem”, where the reward function is allowed to drift mildly across timesteps. This extension is then embedded into the RL setting where, with some additional arguments, we derive the convergence of REINFORCE. In summary, the main contributions of this work are threefold:

- i) We show that the stochastic gradient bandit algorithm will select every arm infinitely often (i.o.) in any bandit problem and with any learning rate. We find it surprising that this strong property emerges from such a simple algorithm, without any explicit hacks to encourage exploration. We obtain a counterpart result in the RL setting, but the bandit case is independently interesting, and also critical for our second contribution:
- ii) In the bandit setting we remove the central assumption of Mei et al. (2024a), that no two arms have the same expected reward, and prove that the stochastic gradient bandit algorithm still converges to an optimal policy. For bandits this assumption is impossible to verify without access to the true reward function (at which point the bandit problem is already solved), but more importantly it renders the extension to RL virtually impossible.
- iii) In RL we provide the first proof that REINFORCE converges with large learning rates in the stochastic setting. This requires the first two contributions: the exploration result is applied directly to RL, and the bandit result is extended to a nonstationary bandit problem that can be embedded into an MDP.

Positioning our work, to our best knowledge, we note that existing convergence results for stochastic policy gradient methods typically suffer from one of the following drawbacks: **(i)** they rely on decaying learning rate schedules for convergence guarantees (Zhang et al., 2020; Ding et al., 2022, 2024; Mei et al., 2023), a requirement not aligned with the constant rates commonly used in practice; **(ii)** results for constant learning rates (Mei et al., 2024b; Klein et al., 2024) provide guarantees only for rates considered impractically small; or **(iii)** they are restricted to the simplest bandit settings (Mei et al., 2024a), limiting their applicability to RL. Filling this gap, our work provides rigorous convergence guarantees for stochastic PG (SPG) with practical learning rates in RL settings, without requiring uniqueness of the optimal policy.

2 Challenges of Non-Unique Solutions

2.1 Non-Uniqueness of Policies in RL

In standard optimization, it is well known that gradient-based algorithms can exhibit non-convergence of their parameters (or iterates) when multiple optimal solutions exist (Absil et al., 2005). To avoid this challenge, existing results for the SPG algorithm in the K -armed bandit setting (Mei et al., 2024b,a) rely on the following assumption, which implies the uniqueness of the globally optimal policy.

Assumption 2.1 (True mean reward has no ties). For all $a, b \in [K]$, if $a \neq b$, then $r(a) \neq r(b)$.

In Assumption 2.1, $[K] := \{1, \dots, K\}$ denotes the set of K arms, and $r(a)$ is the true mean reward for arm $a \in [K]$. Assumption 2.1 implies that there is a unique optimal arm, which we denote $a^* := \arg \max_{a \in [K]} r(a)$. This results in a unique one-hot globally optimal policy π^* with $\pi^*(a^*) = 1$ and $\pi^*(a) = 0$ for all $a \neq a^*$.

However, extending to the RL setting presents the challenge of multiple optimal policies, a scenario which is not prevented by the straightforward extension of Assumption 2.1 to each state, since Assumption 2.1 only constrains immediate rewards. In contrast to bandits, RL involves sequential

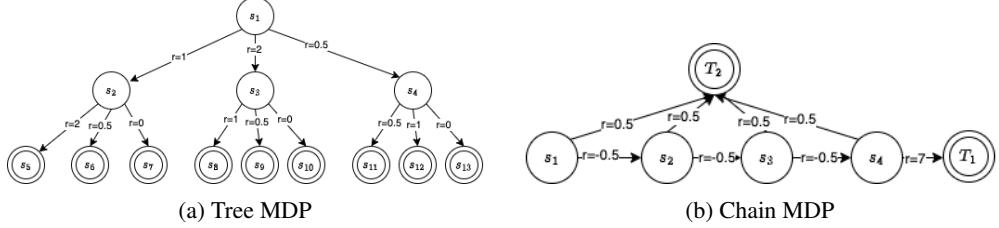


Figure 1: Classical examples of finite-horizon MDPs.

decisions where different action sequences (trajectories) can yield the same maximal cumulative reward. This situation is common in tasks like navigation with alternative optimal paths. As a specific example, consider the tree-structured MDP shown in Fig. 1a (state space $\mathcal{S} = \{s_1, \dots, s_{13}\}$, action space $\mathcal{A} = \{a_1, a_2, a_3\}$). Here, both $s_1 \rightarrow s_2 \rightarrow s_5$ and $s_1 \rightarrow s_3 \rightarrow s_8$ are optimal paths with total reward 3. Because previous bandit convergence analyses (Mei et al., 2024b,a) critically rely on the assumption of a unique optimal policy, they cannot be directly applied to RL problems exhibiting such non-uniqueness.

2.2 SPG Policy Non-Convergence in the Presence of Ties

On the other hand, Mei et al. (2024b, Remark 5.3) conjectured that the SPG algorithm could still achieve convergence even without Assumption 2.1. Their conjecture was based on the idea that SPG has a “self-reinforcing” property, causing the probability of **only one** arm to eventually become dominant and converge to 1, thus resulting in a stationary one-hot optimal policy as $t \rightarrow \infty$. That is, $\pi_t(a^*) \rightarrow 1$ for **only one** optimal arm as $t \rightarrow \infty$, almost surely, **even when multiple optimal arms exist**. If this behavioral property holds, the latter part of the convergence proof can utilize the contradiction-based arguments presented in (Mei et al., 2024b, Theorem 5.1, Claim 2).

Our first major finding, supported by both empirical evidence and theoretical analysis, is that the aforementioned conjecture is incorrect: SPG-like algorithms do not necessarily converge to a **single** policy in the presence of multiple solutions. To demonstrate this, we designed a bandit experiment with two optimal arms (mean 0.2) and one suboptimal arm (mean -0.1). Using the stochastic gradient bandit algorithm (Mei et al., 2024a, Algorithm 1) on a softmax policy ($\theta_0 := \mathbf{0}$, where θ_t are the policy parameters at time t) for 10^5 iterations ($\eta \in \{1, 10\}$), Fig. 2a reveals that, while the total probability of optimal arms converges to 1 ($\sum_{a \in \mathcal{A}^*} \pi_t(a) \rightarrow 1$), the probabilities of individual optimal arms (1 and 2) display non-stationary behavior (e.g. arm 2 fluctuates significantly with $\eta = 1$). We observed analogous behavior in a similar experiment on a tree-structured MDP using REINFORCE (Williams, 1992) ($\eta \in \{0.1, 0.5\}$), as shown in Fig. 2b where optimal action probabilities from state s_1 fail to converge to a unique action. More importantly, we prove the following theorem, rigorously justifying the phenomena observed in simulations.

Proposition 2.2 (Non-Stationary Convergence). *In the bandit setting, where the mean reward has ties, using Algorithm 1 with any $\eta \in \Theta(1)$, for all $a \in \mathcal{A}^*$,*

$$\limsup_t \theta_t(a) = \infty \text{ a.s.} \quad (1)$$

In other words, $(\pi_t)_{t \geq 0}$ does not converge to any one-hot policy.

Proof sketch 1. First, we analyze the dynamics of $(\theta_t(a))_{t \geq 0}$ induced by Algorithm 1. By (Bramson et al., 2004; Davis, 1969, Theorem 1.4), this process must either converge or fluctuate unboundedly. Because it can be shown that $\sum_{t \geq 0} \pi_t(a)(1 - \pi_t(a)) = \infty$, the total variance of the increments $\theta_{t+1}(a) - \theta_t(a)$ is not summable. This implies non-convergence.

2.3 Limitations of Standard Analysis

Global convergence for SPG algorithms is typically established through a two-stage proof: **(i)** establish convergence to a stationary point, and **(ii)** demonstrate (often by contradiction) that the attained stationary point is globally optimal. This methodology originates from seminal work on PG in the exact gradient setting (Agarwal et al., 2020, Theorem 5.1).

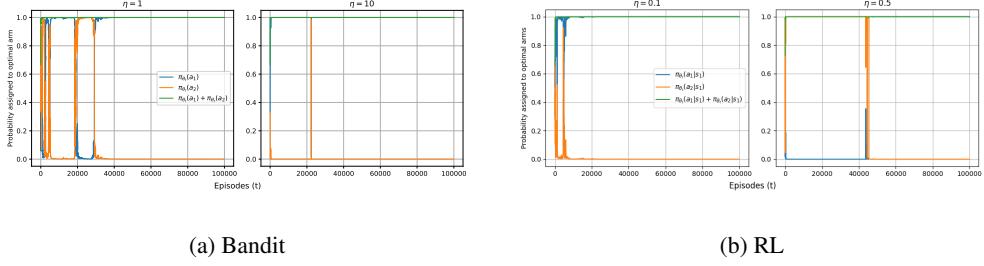


Figure 2: Fig. 2a shows that the total probability of optimal arms converges to 1, but the probabilities of individual optimal arms are non-stationary (i.e. $\pi_t(a_1)$ oscillates). Fig. 2b shows similar non-stationary behavior for optimal actions in an RL setting with multiple optimal trajectories.

Our findings demonstrate two key points: **(i)** ties in trajectory or policy value can exist in RL settings regardless of assumptions on immediate rewards; and **(ii)** the typical two-stage proof strategy for arguing global convergence of SPG cannot be directly extended to RL. However, as shown in the next section and suggested by the above simulations, convergence results **can** be obtained even with ties, but this requires new analysis. This is primarily because our approach needs to carefully reason about the per-timestep expected progress in distinguishing optimal from suboptimal actions despite the presence of these ties. Specifically, we prove that the learned policy eventually converges to assign all probability mass to the optimal set (a form of ‘‘generalized one-hot policy’’), i.e. $\sum_{a \in \mathcal{A}^*} \pi_t(a) \rightarrow 1$ as $t \rightarrow \infty$.

3 An Illustrative Bandit Setting

This section presents new insights into the exploration properties of the SPG algorithm. We first analyze the simplest bandit setting for illustration and then extend the results to RL.

3.1 Stochastic Gradient Bandit

We consider a stochastic multi-armed bandit problem with $K \geq 2$ arms and rewards bounded in $[-R, R]$ (where $R > 0$). At each iteration $t \geq 1$, the learner selects an arm $a_t \in [K] := \{1, \dots, K\}$ and observes a reward r_t sampled from a fixed distribution $P_{a_t} \in \mathcal{M}_1([-R, R])$.¹ The true mean reward for arm $a \in [K]$ is $r(a) := \int_{-R}^R x P_a(dx)$. The set of optimal arms is denoted $\mathcal{A}^* := \arg \max_{a \in [K]} r(a)$.

The learner aims to find a policy $\pi \in \mathcal{M}_1([K])$ that maximizes expected reward. We use the softmax parameterization over \mathbb{R}^K : for $\theta \in \mathbb{R}^K$ and $a \in [K]$,

$$\pi_\theta(a) := \frac{\exp(\theta(a))}{\sum_{b \in [K]} \exp(\theta(b))}. \quad (2)$$

The optimization problem the learner is solving thus has the objective

$$\max_{\theta \in \mathbb{R}^K} \pi_\theta^\top r. \quad (3)$$

We study the stochastic gradient bandit algorithm (Algorithm 1), which performs stochastic gradient ascent on Eq. (3) (Sutton and Barto, 2018; Mei et al., 2024b). Given θ_0 and learning rate $\eta > 0$ the algorithm iteratively updates parameters using the information it receives from single interactions. The stream of parameters generated will be referred to as $(\theta_t)_{t \geq 0}$, and we will use $\pi_t := \pi_{\theta_t}$ for the policy used to select a_{t+1} .

3.2 A Novel Exploration Lemma

We detail the reason why existing results (Mei et al., 2024a) do not generalize, even to bandit settings with reward ties. Mei et al. (2024a, Lemma 2) establishes an exploration property for SPG, showing

¹Where $\mathcal{M}_1(S)$ denotes the collection of probability distributions over the set S .

Algorithm 1 Stochastic gradient bandit algorithm

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1: input  $\theta_0 \in \mathbb{R}^K$ ,  $\eta > 0$ 
2: for  $t \geq 0$  do
3:   Select  $a_{t+1} \sim \pi_t$ , and observe  $r_{t+1} \sim P_{a_{t+1}}$ .
4:    $\theta_{t+1}(a_{t+1}) \leftarrow \theta_t(a_{t+1}) + \eta(1 - \pi_t(a_{t+1}))r_{t+1}$ .
5:   for  $a \in [K]$ ,  $a \neq a_{t+1}$  do
6:      $\theta_{t+1}(a) \leftarrow \theta_t(a) - \eta\pi_t(a)r_{t+1}$ .
7:   end for
8: end for

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that at least two **distinct** arms are sampled i.o. Their subsequent convergence proof (Mei et al., 2024a, Theorem 2) relies on the argument that at least one of these i.o. sampled actions must be optimal. However, in the presence of reward ties, it is possible for two actions to share the same reward value (the sub-optimal action's interval from (Mei et al., 2024a, Eq. (15)) no longer exists). Consequently, the arguments that construct a contradiction to show “at least one of these i.o. sampled actions must be optimal” are no longer valid.

Given the failure of existing approaches with reward ties, new analytical results are required for convergence proofs, even in the bandit setting. Our second key finding is a generalized exploration property for SPG: we establish that despite reward ties, **every arm** is sampled i.o. To formalize this, we define $N_t(a)$ as the number of times action $a \in [K]$ has been sampled up to iteration $t \geq 1$, i.e.

$$N_t(a) := \sum_{s=1}^t \mathbb{I}\{a_s = a\}. \quad (4)$$

The asymptotic count is $N_\infty(a) := \lim_{t \rightarrow \infty} N_t(a)$, which is either finite or infinite. If $N_\infty(a) < \infty$, action a is only sampled finitely many times; if $N_\infty(a) = \infty$, action a is sampled i.o.

Lemma 3.1 (Bandit Exploration). *Using Algorithm 1 with any constant learning rate $\eta \in \Theta(1)$, every arm is almost surely played infinitely often. That is, $\forall a \in [K] : N_\infty(a) = \infty$ almost surely.*

Proof sketch 2. For any arm $a' \in [K]$ such that $N_\infty(a') < \infty$, the Extended Borel-Cantelli (Breiman, 1992) Lemma implies $\sum_{t=0}^\infty \pi_t(a') < \infty$. Since such an arm is sampled only finitely many times, its parameter $\theta_t(a')$ remains bounded, $\sup_t |\theta_t(a')| < \infty$, and its probability converges to zero: $\lim_{t \rightarrow \infty} \pi_t(a') = 0$. Without loss of generality, let $a \in [K]$ be an arm with $N_\infty(a) < \infty$. The condition $\lim_{t \rightarrow \infty} \pi_t(a) = 0$ requires that some parameter grows unboundedly, i.e. $\lim_{t \rightarrow \infty} \max_{a' \in [K]} \theta_t(a') = \infty$. To preserve the total probability mass, this necessitates that some parameter must diverge to negative infinity: $\lim_{t \rightarrow \infty} \min_{a' \in [K]} \theta_t(a') = -\infty$. Thus, there exists at least one arm $b \in [K]$ such that $\liminf_{t \rightarrow \infty} \theta_t(b) = -\infty$. Furthermore, since the sum of probabilities for finitely sampled arms is finite, any arm b with $\liminf_{t \rightarrow \infty} \theta_t(b) = -\infty$ must be sampled infinitely often ($N_\infty(b) = \infty$).

We use these properties of arms a (finitely sampled, bounded parameter) and b (infinitely sampled, parameter unbounded below) to construct a proof by contradiction. The fact that arm b is sampled infinitely often despite its parameter repeatedly dropping to arbitrarily low values implies that $\theta_t(b)$ must periodically increase to become larger than $\theta_t(a)$ (and other bounded parameters) infinitely often. Consider the event $C_t := \{\theta_t(b) < \theta_t(a), a_t = b\}$. We first show that if $\theta_t(b) \leq \theta_t(a)$ and the parameter update causes $\theta_{t+1}(b) > \theta_{t+1}(a)$, this implies $a_t = b$. We then prove that the event C_t occurs only a finite number of times. However, for arm b to be sampled infinitely often ($N_\infty(b) = \infty$) while $\liminf \theta_t(b) = -\infty$ and $\theta_t(a)$ is bounded, it must be sampled infinitely often during periods when $\theta_t(b) < \theta_t(a)$. This contradicts the finding that C_t occurs only finitely often, proving our initial assumption ($N_\infty(a) < \infty$ for some arm a) is false.

3.3 Convergence Without the Assumption of Unique Rewards

Our new result about the exploration of SPG in the bandit setting, Lemma 3.1, allows us to remove an assumption necessary for the results of prior work (Mei et al., 2024a,b), namely that there are no ties in the true mean rewards of the arms (Assumption 2.1). However, this requires new analysis beyond the exploration proof. In this section we sketch out the steps used to show our central result in the bandit setting: that Algorithm 1 converges almost surely regardless of the learning rate.

Theorem 3.2 (Convergence in Bandits). *In the bandit setting of Section 3.1 without Assumption 2.1, Algorithm 1 with any $\eta \in \Theta(1)$ almost surely converges to playing optimal arms,*

$$\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \pi_t(a) = 1 \text{ a.s.} \quad (5)$$

The proof of this theorem breaks into two propositions, the first of which being that the sum of parameters of optimal arms tends to infinity (excluding the trivial case where all arms are equally good and Section 3.1 holds vacuously).

Proposition 3.3 (Infinite Optimal Parameters). *If $\mathcal{A}^* \neq [K]$ then $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty$ a.s.*

The second proposition states that all finite arms individually have their parameters diverge to negative infinity.

Proposition 3.4 (Negative Infinite Suboptimal Parameters). *For every suboptimal arm $b \in [K] \setminus \mathcal{A}^*$, $\lim_{t \rightarrow \infty} \theta_t(b) = -\infty$ a.s.*

Equipped with these two propositions, the proof of Theorem 3.2 becomes straightforward enough that we need not resort to a proof sketch:

Proof of Theorem 3.2. If $\mathcal{A}^* = [K]$ then $\sum_{a \in \mathcal{A}^*} \pi_t(a) = 1$ for all $t \geq 0$ and the result holds vacuously. Henceforth suppose $\mathcal{A}^* \neq [K]$. We have that $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \pi_t(a) = 1 - \lim_{t \rightarrow \infty} \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b)$, so it suffices to show that, for all $b \in [K] \setminus \mathcal{A}^*$, $\lim_{t \rightarrow \infty} \pi_t(b) = 0$. To this end fix $b \in [K] \setminus \mathcal{A}^*$. We have the following bound from expanding the definition of π_t :

$$\lim_{t \rightarrow \infty} \pi_t(b) = \lim_{t \rightarrow \infty} \frac{\exp(\theta_t(b))}{\sum_{a \in [K]} \exp(\theta_t(a))} \quad (\text{Eq. (2)}) \quad (6)$$

$$\leq \lim_{t \rightarrow \infty} \frac{\exp(\theta_t(b))}{\sum_{a \in \mathcal{A}^*} \theta_t(a)} \quad (\exp(x) \geq x, \mathcal{A}^* \subset [K]) \quad (7)$$

$$= \frac{\lim_{t \rightarrow \infty} \exp(\theta_t(b))}{\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \theta_t(a)}. \quad (8)$$

Proposition 3.4 implies that the upper limit in Eq. (8) approaches 0 and Proposition 3.3 implies that the lower limit goes to infinity. Thus $\lim_{t \rightarrow \infty} \pi_t(b) = 0$, concluding the proof. \square

The proofs of Propositions 3.3 and 3.4 are long and technical, and we refer the reader to the appendix for the details.

4 Reinforcement Learning

The results in RL depend on the results of Section 3, but in order to apply them we will need to port them to a slightly generalized bandit problem. We describe the necessary modifications in the following subsection, before proceeding to MDPs.

4.1 Nonstationary Bandit Setting

We still consider a K -armed bandit, with $K \geq 2$ and rewards in $[-R, R]$. The interaction between the learner and the environment is much the same as in Section 3.1, with the exception that now the reward distributions are allowed to change across timesteps. That is, we change out the distribution $P_a \in \mathcal{M}_1([-R, R])$ of rewards given that arm a is played with a sequence of such distributions $(P_a^t)_{t \geq 1}$, and the reward at each iteration $t \geq 1$ is sampled from $P_{a_t}^t \in \mathcal{M}_1([-R, R])$; we also allow the expected rewards given that an arm is played to vary over time, so $r(a)$ becomes $(r^t(a))_{t \geq 1}$, and we have $\mathbb{E}[r_t | a_t = a] = r^t(a)$.

However, we constrain the setting in two ways. First, we suppose that there exists a filtration $(\mathcal{F}_t)_{t \geq 0}$ such that P^t, r^t are \mathcal{F}_{t-1} -measurable and a_t, r_t are \mathcal{F}_t -measurable. Intuitively, \mathcal{F}_t contains the information available to the learner at iteration t , and this assumption means that the reward distributions (and thus their means) may only depend on the arms played and rewards observed

up to the current timestep, as well as additional sources of randomness that are independent of the future. The second constraint on the environment is that we assume the existence of a “true” mean reward vector $r \in [-R, R]^K$, and suppose that there exists some random timestep τ such that, for all $t \geq \tau$ and all $a \in [K]$, $|r(a) - r^t(a)| \leq \Delta/3$, where $\Delta := \min_{a,b \in [K]: r(a) \neq r(b)} |r(a) - r(b)|$ is the minimum nonzero gap in the “true” mean reward between any two arms. This says that eventually the expected reward of playing arm a will settle down to a neighbourhood of $r(a)$, and in particular that the arms in $\mathcal{A}^* := \arg \max_{a \in [K]} r(a)$ have the highest expected reward after iteration τ . Given these modifications to the bandit setting, we can extend the results of Section 3 with minimal changes. The algorithm stays exactly the same, with the only modification to Algorithm 1 being that, at line 3, $r_{t+1} \sim P_{a_{t+1}}$ becomes $r_{t+1} \sim P_{a_{t+1}}^{t+1}$.

After extending all the bandit results to the nonstationary bandit setting, we can finally apply them for a result in RL.

4.2 Reinforcement Learning Setting

We consider a finite-horizon MDP, defined by the tuple $\mathcal{M} = (\mathcal{H}, \mathcal{S}, \mathcal{A}, \{r_h\}_{h=0}^{H-1}, \{P_h\}_{h=0}^{H-1}, \rho)$, where $\mathcal{H} = \{0, 1, \dots, H-1\}$ is the index set of timesteps in an episode; $\mathcal{S} = \mathcal{S}_0 \cup \dots \cup \mathcal{S}_{H-1}$ and $\mathcal{A} = \mathcal{A}_0 \cup \dots \cup \mathcal{A}_{H-1}$ are finite state and action spaces, respectively, with \mathcal{S}_h ($\mathcal{A}_h = \cup_{s \in \mathcal{S}_h} \mathcal{A}_s$) being the sets of possible states and (actions) at step $h \in \mathcal{H}$, and \mathcal{A}_s is the set of possible actions from state s ; $r_h : \mathcal{S}_h \times \mathcal{A}_h \rightarrow [-R, R]$ is a reward function that is bounded by $R > 0$; $P_h : \mathcal{S}_h \times \mathcal{A}_h \rightarrow \mathcal{M}_1(\mathcal{S}_{h+1})$ is the transition function; and $\rho : \mathcal{S}_0 \rightarrow \mathcal{M}_1(\mathcal{S}_0)$ is the initial state distribution. We denote $\pi := (\pi^h)_{h=0}^{H-1}$ as a time-dependent policy where $\pi^h : \mathcal{S}_h \rightarrow \mathcal{M}_1(\mathcal{A}_h)$ is the policy in the horizon h . An episode proceeds under the following protocol. At the beginning of the episode, the learner selects a non-stationary policy π . The episode then evolves through $s_0 \sim \rho$ and $a_h \sim \pi^h(\cdot | s_h)$, $s_{h+1} \sim p_h(\cdot | s_h, a_h)$, $r_h = r_h(a_h, s_h)$ for all $h \in \mathcal{H}$. We define the trajectory $\tau := (s_0, a_0, r_0, s_1, \dots, s_{H-1}, a_{H-1}, r_{H-1})$. Therefore, the probability of a given trajectory τ is

$$\Pr(\tau) = \rho(s_0) \pi^0(a_0 | s_0) p_0(s_1 | s_0, a_0) \dots \pi_{H-1}(a_{H-1} | s_{H-1}) \quad (9)$$

We also define the value functions and action-value functions for $h \in \mathcal{H}$

$$V_h^\pi(s) := \mathbb{E}^\pi \left[\sum_{h'=h}^{H-1} r_{h'} \middle| s_h = s \right] \quad (10)$$

$$Q_h^\pi(s, a) := \mathbb{E}^\pi \left[\sum_{h'=h}^{H-1} r_{h'} \middle| s_h = s, a_h = a \right] \quad (11)$$

The goal is to find a time-dependent policy π^* that maximizes the state-value function at time 0, i.e. $V_0^\pi(\rho) := \mathbb{E}_{s \sim \rho}[V_0^\pi(s)]$:

$$\pi^* \in \arg \max_{\pi} V_0^\pi(\rho). \quad (12)$$

We also define optimal state and state-action value function, $V_h^* := V_h^{\pi^*}$ and $Q_h^* := Q_h^{\pi^*}$. In this paper, we focus on softmax parameterized policies. Specifically, we parameterize each π^h by θ^h for all $h \in \mathcal{H}$ by

$$\pi_{\theta^h}^h(a | s) := \frac{\exp(\theta^h(s, a))}{\sum_{a'} \exp(\theta^h(s, a'))} \quad (13)$$

where $\theta^h \in \mathbb{R}^{\mathbb{A}_h}$ with $\mathbb{A}_h := \sum_{s \in \mathcal{S}_h} |\mathcal{A}_s|$ for all $h \in \mathcal{H}$. To improve the readability, we will sometimes write π_t in place of π_{θ_t} and π_t^h in place of $\pi_{\theta_t}^h$. The true gradient of the Eq. (12) is

$$\frac{\partial V_0^\pi(s)}{\partial \theta_t^h(s, a)} = \mathbb{E}^\pi \left[\frac{\partial}{\partial \theta_t^h(s, a)} \log \pi_t(a_h | s_h) \sum_{h=0}^{H-1} r_h \right] = \mathbb{E}^\pi \left[(\mathbb{I}[a_h = a] - \pi_t^h(a | s)) \sum_{h=0}^{H-1} r_h \right] \quad (14)$$

where $\mathbb{I}[a_h = a]$ is the indicator function of whether action a is played in the horizon h , for all $s \in \mathcal{S}_h, a \in \mathcal{A}_h, h \in \mathcal{H}$. Since we are in the stochastic setting, we will use REINFORCE estimator to estimate the gradient and update the parameters

$$\frac{\hat{\partial} V_0^\pi(s)}{\partial \theta_h(s, a)} = \left(\sum_{h'=h}^{H-1} r_{h'} \right) \left(\mathbb{I}[a_h = a] - \pi_t^h(a | s) \right) \quad (15)$$

Algorithm 2 REINFORCE

```

1: for each episode do
2:   Sample a trajectory  $\tau$  using  $\rho, \{\pi_{\theta_h}\}_{h=1}^{H-1}, \{P_h\}_{h=1}^{H-1}$ 
3:   for all  $a \in |\mathcal{A}|, s \in |\mathcal{S}|$  do
4:     Use Eq. (15) to update  $\theta(s, a)$ 
5:   end for
6: end for

```

The REINFORCE algorithm is shown in Algorithm 2.

We first show an exploration result, the counterpart to the exploration result shown above in the bandit setting, before sketching the proof of our main theorem in RL.

4.3 RL Exploration Lemma

Lemma 4.1. *Running REINFORCE with any $\eta \in \Theta(1)$ in a finite-horizon MDP \mathcal{M} , for all $h \in \mathcal{H}$, for all reachable $s \in \mathcal{S}_h$ and for all $a \in \mathcal{A}_s$ we have, almost surely, that every reachable state action pair will be visited i.o, i.e $N_\infty(s, a) = \infty$.*

Proof sketch 3. First, we show that for all horizon $h \in \mathcal{H}$, if $s \in \mathcal{S}_h$ is reachable and played i.o, then all actions $a \in \mathcal{S}$ are also played i.o by Lemma 3.1. Next, we use induction to show that for all horizon $h \in \mathcal{H}$, if $s \in \mathcal{S}_h$ is reachable visited i.o, then $s' \in \mathcal{S}_{h+1}$ is also visited i.o. Therefore, for all $h \in \mathcal{H}$, all reachable state-action pairs $(s, a) \in \mathcal{S} \times \mathcal{A}$ will be played i.o.

4.4 Convergence in finite-horizon MDP

Theorem 4.2 (Convergence in RL). *For the MDP defined as above, using REINFORCE with constant learning rate $\eta \in \Theta(1)$, we have, almost surely, for all $s \in \mathcal{S}_0$, $V_0^{\pi_t}(s) \rightarrow V_0^*(s)$ as $t \rightarrow \infty$*

Proof sketch 4. We show the convergence theorem using the backward induction. Suppose for all horizon $h \in \{h', \dots, H-1\}$, we have $\sum_{a \in \mathcal{A}_s^*} \pi_t^h(a|s) \rightarrow 1$ for all $s \in \mathcal{S}_h$, we want to show that $\sum_{a \in \mathcal{A}_s^*} \pi_t^{h-1}(a|s) \rightarrow 1$ for all $s \in \mathcal{S}_{h-1}$. Since $\sum_{a \in \mathcal{A}_s^*} \pi_t^h(a|s) \rightarrow 1$ for all $s \in \mathcal{S}_h$, where $h \in \{h', \dots, H-1\}$, we know that there exists time step τ s.t $V_h^*(s) - V_h^{\pi_t}(s) \leq \frac{\delta}{3}$, where δ is the minimum non-zero gap between two Q-values. For all $a \in \mathcal{A}_{s_{h-1}}$, there exists a minimum gap of $\frac{\delta}{3}$ in the Q-value. Therefore, applying the bandit convergence result, we know that $\sum_{a \in \mathcal{A}_s^*} \pi_t^{h-1}(a|s) \rightarrow 1$ as $t \rightarrow \infty$ for all $s \in \mathcal{S}_{h-1}$. Recursively, we know that for all $s \in \mathcal{S}_0$, $\sum_{a \in \mathcal{A}_s^*} \pi_{\theta_t}^0(a|s) \rightarrow 1$ as $t \rightarrow \infty$.

We also provide the statement and proof of convergence rate in the appendix (Theorem E.3).

4.5 Simulations

We conduct several experiments to illustrate the convergence behavior of REINFORCE algorithm in the finite-horizon setting. Experiments are performed using a chain MDP (Fig. 1b) with state space $\mathcal{S} = \{s_0, \dots, s_3, T_1, T_2\}$, where T_1 and T_2 are terminal states, and action space $\mathcal{A} = \{a_0, a_1\}$. Taking action a_0 in any state yields a mean reward of 0.5 and transitions to a terminal state T_1 . Taking action a_1 in state s_i ($i \in \{0, 1, 2\}$) yields a mean reward of -0.5 and transitions to state s_{i+1} . In state s_3 , action a_1 yields a mean reward of 7 and transitions to a terminal state T_2 . The policy is parameterized using a softmax function, and parameters are initialized to $\mathbf{0} \in \mathbb{R}^{|\mathcal{S}| \times |\mathcal{A}|}$. For each learning rate η , the REINFORCE algorithm is run for 10^5 episodes across 30 seeds. Performance is evaluated by measuring the average suboptimality gap from the initial state distribution ρ , defined as $V_0^*(\rho) - V_0^{\pi_t}(\rho)$, over the 30 seeds. Our first experiment (Fig. 3a) demonstrates the benefits of using a large learning rate. Previous convergence analysis of REINFORCE (Theorem 4.1 Klein et al., 2024) relies on small constant learning rates, which can significantly impede practical training speed. For instance, the analysis in Klein et al. (2024) guarantees convergence with $\eta = \frac{1}{5H^2R\sqrt{T}}$, where T is the number of training episodes. In our environment ($H = 4, R = 7, T = 10^5$), this corresponds to an extremely small learning rate $\eta \approx 10^{-7}$. Therefore, we evaluated REINFORCE algorithm with larger learning rates $\eta \in \{0.00001, 0.001, 0.1\}$. Fig. 3a shows that the suboptimality gap remains

nearly constant for $\eta = 0.00001$, whereas it decreases substantially faster as η increases from 0.001 to 0.1. This demonstrates the practical benefit of employing larger learning rates for accelerated convergence, supported by our theoretical guarantees. We further explore the effect of even larger learning rates $\eta \in \{0.5, 1, 2\}$, presented in Fig. 3b. These rates, while potentially accelerating learning if updates are favorable, generally slow down convergence compared to the moderately large rates. The suboptimality curves exhibit more abrupt changes and show less consistent improvement over episodes. The large shaded regions indicate significantly higher variance with these very large learning rates. This suggests that large steps can easily push parameters away from optimal configurations, leading to prolonged exploration of suboptimal regions until a corrective update is sampled. Finally, Fig. 4 illustrates the evolution of the learned policy for optimal actions at each horizon. For all learning rates, we observe, on average, that the probability of selecting optimal actions converges first for the last horizon, then for the second-to-last, and so on, proceeding backward through the horizon. This backward convergence pattern in policy probabilities is consistent with our proof strategy for the convergence of the REINFORCE algorithm, which relies on a backward induction approach. We also extend our experiments to demonstrate the relationship between the algorithm's performance and different learning rates. Details on the experimental setups and results can be found in the section Appendix F. Overall, we consistently find a "bowl-shaped" relationship between the learning rate and performance, meaning both excessively small and excessively large learning rates lead to high suboptimality, while middling values achieve the smallest suboptimality. The specific shape and optimal point of this bowl vary significantly with the environment's structure.

Remark 4.3. It is worth noting that not all environments exhibit this specific backward convergence pattern in learning optimal policy. However, the presence or absence of this pattern in empirical observations does not invalidate our main theoretical result.

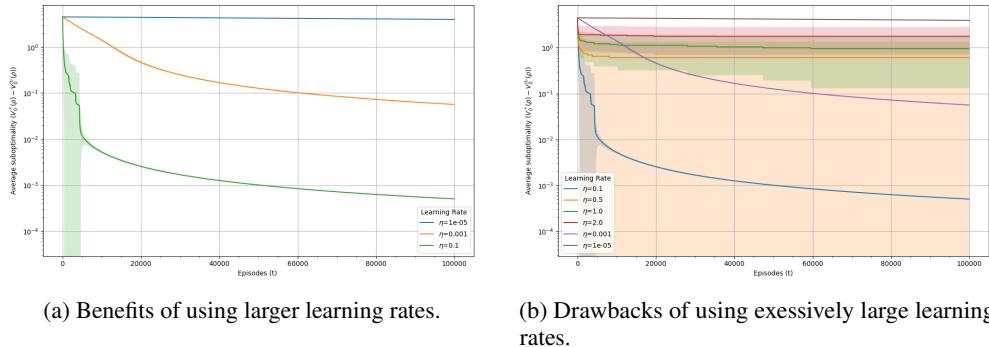


Figure 3: Fig. 3a shows that using a larger learning rate can improve the performance of REINFORCE, while Fig. 3b shows that excessively large learning rates have substantial variance, which can slow down the convergence rate.

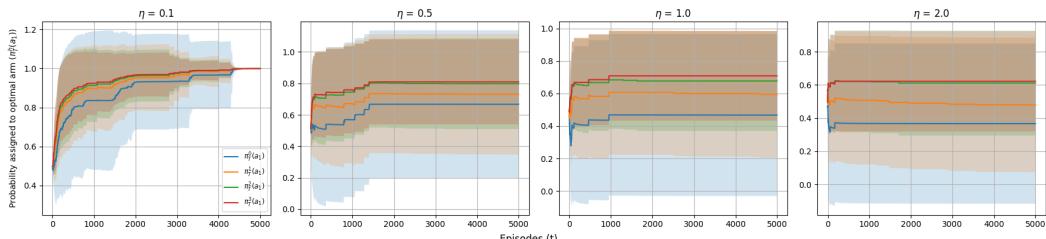


Figure 4: These figures show the convergence rate of the optimal policy in each horizon for different learning rates. In particular, we observe that the optimal policy of the last horizon will converge first, then the second-to-last one until the first horizon. This observation aligns with our analysis.

5 Conclusions and Future work

This work enhances our understanding of the convergence properties of the widely used REINFORCE algorithm. Our novel proof offers deeper insights into the exploration effects of stochastic gradient methods and raises new research questions. Notably, recent findings by Mei et al. (2024a) indicate a convergence rate of $O(\log(t)/t)$ for stochastic gradient bandit algorithms. This has a gap with the established $O(1/t)$ lower bound for SPG Mei et al. (2021), suggesting a potential for accelerated convergence in bandit settings and, by extension, in RL setting. As demonstrated in Fig. 3b, REINFORCE with excessively large learning rates exhibits high variance, impeding convergence. Future work could explore optimal learning rate schedules to harness the initial benefits of larger rates while subsequently mitigating variance. Other promising directions include extending the convergence result for REINFORCE to function approximation setting.

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In this appendix we will deal repeatedly with almost sure events, i.e. events that occur with probability 1. We typically mention this throughout the proofs, except for in one important case where “a.s.” is omitted to reduce clutter: whenever statements involving conditional expectations (by extension conditional probabilities, variances) do not have an explicit probabilistic quantification, they are understood to hold almost surely. Of course, this is the only possible interpretation for such statements, since conditional expectations are only defined up to a set of measure 0.

A Technical Tools

We begin with some fundamental results from probability theory. The first is a generalization of the Borel-Cantelli Lemma.

Lemma A.1 (Extended Borel-Cantelli Lemma, Corollary 5.29 of Breiman (1992)). *Given a filtration $(\mathcal{F}_t)_{t \geq 0}$ and a sequence of events $(A_t)_{t \geq 0}$ with $A_t \in \mathcal{F}_t$ for all $t \geq 0$,*

$$\sum_{t \geq 0} \mathbb{I}[A_t] = \infty \iff \sum_{t \geq 0} \mathbb{P}(A_t | \mathcal{F}_{t-1}) = \infty. \quad (16)$$

That is, $(A_t)_{t \geq 0}$ occurs i.o. if and only if $\sum_{t \geq 0} \mathbb{P}(A_t | \mathcal{F}_{t-1})$ is infinite, up to a set of measure zero.

Our analysis relies critically and repeatedly on the celebrated inequality of Freedman. The version we will use is similar to the one stated by Mei et al. (2024a,b). Since we require a general filtration, we include the original statement by Freedman below in Lemma A.2, followed by the statement and derivation of the form most convenient to us in Lemma A.3. Whenever we mention “Freedman’s inequality” elsewhere in this work it shall refer to the latter.

Lemma A.2 ((Original) Freedman’s Inequality, Theorem 1.6 of Freedman (1975)). *Given a filtered probability space with filtration $(\mathcal{F}_t)_{t \geq 0}$, an adapted sequence of random variables $(X_t)_{t \geq 1}$, and constants $a, b > 0$, if $\forall t \geq 1 : \mathbb{E}[X_t | \mathcal{F}_{t-1}] = 0$ and $|X_t| \leq 1$ then*

$$\mathbb{P}\left(\exists t \geq 1 : \sum_{i \in [t]} X_i \geq a, \sum_{i \in [t]} \text{Var}[X_i | \mathcal{F}_{i-1}] \leq b\right) \leq \exp\left(\frac{-a^2}{2(a+b)}\right). \quad (17)$$

Lemma A.3 (Freedman’s Inequality). *Let $(X_t)_{t \geq 1}$ be a random sequence adapted to the filtration $(\mathcal{F}_t)_{t \geq 0}$, $B \geq 0$ be a constant such that $\forall t \geq 0 : |X_t| \leq B$, and denote $V_t := \sum_{i \in [t]} \text{Var}[X_i | \mathcal{F}_{i-1}]$. For any $\delta \in (0, 1]$, it holds with probability $1 - \delta$ that*

$$\forall t \geq 1 : \left| \sum_{i \in [t]} X_i - \mathbb{E}[X_i | \mathcal{F}_{i-1}] \right| \leq 20\sqrt{V_t + 4B^2 + 1} \log\left(\frac{V_t + 2}{\delta}\right). \quad (18)$$

Remark A.4. The derivation of Lemma A.3 closely follows the proof of Theorem C.3 of Mei et al. (2024b). We aimed for a simple bound rather than a tight one.

Proof. Fix $\epsilon \in (0, 1)$, and let $S_t := \sum_{i \in [t]} X_i$ and $V_t := \sum_{i \in [t]} \text{Var}[X_i | \mathcal{F}_{i-1}]$. First we will suppose that $\mathbb{E}[X_t | \mathcal{F}_{t-1}] = 0$ and $|X_t| \leq 1$ for all $t \geq 1$, and show that

$$\mathbb{P}\left(\exists t \geq 1 : S_t \geq 10\sqrt{V_t + 1} \log\left(\frac{V_t + 2}{\epsilon}\right)\right) \leq \epsilon. \quad (19)$$

For $x \geq 1$ let $g(x) := 3\log((x+2)^2/\epsilon)$, and we have

$$g(x) + \sqrt{g(x)x} \quad (20)$$

$$\leq 6\log((x+2)/\epsilon) + \sqrt{3x}\sqrt{\log((x+2)^2/\epsilon)} \quad ((x+2)^2/\epsilon \leq (x+2)^2/\epsilon^2) \quad (21)$$

$$\leq 6\log((x+2)/\epsilon) + \sqrt{3x}\log((x+2)^2/\epsilon) \quad (\log((x+2)^2/\epsilon) \geq \log(4) \geq 1) \quad (22)$$

$$\leq 6\log((x+2)/\epsilon) + 2\sqrt{3x}\log((x+2)/\epsilon) \quad ((x+2)^2/\epsilon \leq (x+2)^2/\epsilon^2) \quad (23)$$

$$\leq 6\log((x+2)/\epsilon) + 4\sqrt{x+1}\log((x+2)/\epsilon) \quad (24)$$

$$\leq 10\sqrt{x+1}\log((x+2)/\epsilon). \quad (25)$$

Setting $x := V_t$ in Eq. (25) yields

$$\mathbb{P}\left(\exists t \geq 1 : S_t \geq 10\sqrt{V_t + 1} \log\left(\frac{V_t + 2}{\epsilon}\right)\right) \quad (26)$$

$$\leq \mathbb{P}\left(\exists t \geq 1 : S_t \geq g(V_t) + \sqrt{g(V_t)V_t}\right) \quad (27)$$

$$= \sum_{i \geq 0} \mathbb{P}(\exists t \geq 1 : S_t \geq g(V_t) + \sqrt{g(V_t)V_t}, i \leq V_t < i + 1) \quad (28)$$

$$= \sum_{i \geq 0} \mathbb{P}(\exists t \geq 1 : S_t \geq g(i) + \sqrt{g(i)i}, V_t \leq i + 1) \quad (g, \sqrt{\cdot} \text{ are increasing}) \quad (29)$$

$$\leq \sum_{i \geq 0} \exp\left(-\frac{(g(i) + \sqrt{g(i)i})^2}{2(g(i) + \sqrt{g(i)i} + i + 1)}\right).$$

(Lemma A.2 with $a := g(i) + \sqrt{g(i)i}$ and $b := i + 1$) (30)

To control the term appearing in the exp above, we will use the following inequality, which holds for $u \geq 2$ and $i \geq 0$:

$$\frac{(u + \sqrt{ui})^2}{2(u + \sqrt{ui} + i + 1)} = \frac{u(u + 2\sqrt{ui} + i)}{2(u + \sqrt{ui} + i + 1)} \quad (31)$$

$$= \frac{u}{3} \cdot \frac{2u + 6\sqrt{ui} + 3i + u}{2u + 2\sqrt{ui} + 2i + 2} \quad (32)$$

$$\geq u/3. \quad (u \geq 2) \quad (33)$$

Since $g(i) \geq 3\log(4) \geq 2$, we can combine the above two displays by setting $u := g(i)$ and conclude

$$\mathbb{P}\left(\exists t \geq 1 : S_t \geq 10\sqrt{V_t + 1} \log\left(\frac{V_t + 2}{\epsilon}\right)\right) \leq \sum_{i \geq 0} \exp(-g(i)/3) \quad (34)$$

$$= \epsilon \sum_{i \geq 0} \frac{1}{(i + 2)^2} \quad (35)$$

$$= \epsilon \sum_{i \geq 2} i^{-2} \quad (36)$$

$$= \epsilon(\pi^2/6 - 1) \quad (\sum_{i \geq 1} i^{-2} = \pi^2/6) \quad (37)$$

$$\leq \epsilon. \quad (38)$$

We are finished showing Eq. (19). We can apply this result to both $(X_t)_{t \geq 0}$ and $(-X_t)_{t \geq 0}$, setting $\epsilon := \delta/2$ in each application, whence a union bound guarantees that, with probability at least $1 - \delta$,

$$\forall t \geq 1 : \sum_{i \in [t]} |X_i| < 10\sqrt{V_t + 1} \log\left(\frac{V_t + 2}{\delta/2}\right) \quad (39)$$

$$\leq 10\sqrt{V_t + 1} \left(\log\left(\frac{V_t + 2}{\delta}\right) + \log(2) \right) \quad (40)$$

$$\leq 20\sqrt{V_t + 1} \log\left(\frac{V_t + 2}{\delta}\right). \quad (41)$$

Given a random sequence $(X_t)_{t \geq 1}$ that satisfies $|X_t| \leq B$ for some $B \geq 1$, we can apply Eq. (41) to the sequence $(X_t/B)_{t \geq 1}$:

$$\forall t \geq 1 : \sum_{i \in [t]} |X_i/B| < 20\sqrt{V_t/B^2 + 1} \log\left(\frac{V_t/B^2 + 2}{\delta}\right) \quad (42)$$

$$\leq 20\sqrt{V_t/B^2 + 1} \log\left(\frac{V_t + 2}{\delta}\right) \quad (V_t/B^2 \leq V_t) \quad (43)$$

$$\Rightarrow \sum_{i \in [t]} |X_i| \leq 20\sqrt{V_t + B^2} \log\left(\frac{V_t + 2}{\delta}\right), \quad (44)$$

with probability $1 - \delta$. Combining Eq. (41), which holds for $|X_t| \leq 1$, and Eq. (44), which holds for $|X_t| \leq B$ where $B \geq 1$, and upper bounding $\max(B^2, 1) \leq B^2 + 1$, we can remove the requirement that $B \geq 1$ and conclude that, if $\mathbb{E}[X_t | \mathcal{F}_{t-1}] = 0$ and $|X_t| \leq B$ for all $t \geq 1$, with probability $1 - \delta$

$$\forall t \geq 1 : \sum_{i \in [t]} |X_i| \leq 20\sqrt{V_t + B^2 + 1} \log\left(\frac{V_t + 2}{\delta}\right). \quad (45)$$

To remove the assumption that $\forall t \geq 1 : \mathbb{E}[X_t | \mathcal{F}_{t-1}] = 0$ and get the desired result, we apply Eq. (45) to $(X_t - \mathbb{E}[X_t | \mathcal{F}_{t-1}])_{t \geq 1}$, noting that if $|X_t| \leq B$ then $|X_t - \mathbb{E}[X_t | \mathcal{F}_{t-1}]| \leq 2B$. \square

The following result applies Freedman's Inequality to a sequence of bounded, and eventually (conditionally) self-bounded, random variables. It says that if the conditional expectations are not summable then the variables themselves will not be summable. We expect that the result is folklore, but cannot find a reference.

Lemma A.5 (Freedman Divergence Trick). *Let $(X_t)_{t \geq 1}$ be a random sequence adapted to the filtration $(\mathcal{F}_t)_{t \geq 0}$ and $B \geq 0$ be a constant such that $\forall t \geq 0 : |X_t| \leq B$. Suppose $\sum_{t \geq 1} \mathbb{E}[X_t | \mathcal{F}_{t-1}] = \infty$ and, for some random (a.s. finite) index $\tau \geq 1$ and constant $C \geq 0$, for all $t \geq \tau$, $\text{Var}[X_t | \mathcal{F}_{t-1}] \leq C\mathbb{E}[X_t | \mathcal{F}_{t-1}]$. Then $\sum_{t \geq 1} X_t = \infty$ a.s.*

Remark A.6. Note that the result does not require τ to be a stopping time.

Proof. For $t \geq 0$ let $S_t := \sum_{i \in [t]} X_i$, $\bar{S}_t := \sum_{i \in [t]} \mathbb{E}[X_i | \mathcal{F}_{i-1}]$, and $V_t := \sum_{i \in [t]} \text{Var}[X_i | \mathcal{F}_{i-1}]$. For any $\delta \in (0, 1]$, we can apply Freedman's Inequality (Lemma A.3) to $(X_t)_{t \geq 1}$. This gives that, with probability $1 - \delta$, for any $t \geq \tau$,

$$S_t \geq \bar{S}_t - 20\sqrt{V_t + 4B^2 + 1} \log\left(\frac{V_t + 2}{\delta}\right) \quad (46)$$

$$= \bar{S}_t - \bar{S}_\tau + \bar{S}_\tau - 20\sqrt{V_t - V_\tau + V_\tau + 4B^2 + 1} \log\left(\frac{V_t - V_\tau + V_\tau + 2}{\delta}\right) \quad (47)$$

$$\geq \bar{S}_t - \bar{S}_\tau - \tau B - 20\sqrt{V_t - V_\tau + (4 + \tau)B^2 + 1} \log\left(\frac{V_t - V_\tau + \tau B^2 + 2}{\delta}\right) \quad (|X_t| \leq B) \quad (48)$$

$$\geq \bar{S}_t - \bar{S}_\tau - \tau B - 20\sqrt{C(\bar{S}_t - \bar{S}_\tau) + (4 + \tau)B^2 + 1} \log\left(\frac{C(\bar{S}_t - \bar{S}_\tau) + \tau B^2 + 2}{\delta}\right). \quad (\text{Var}[X_i | \mathcal{F}_{i-1}] \leq C\mathbb{E}[X_i | \mathcal{F}_{i-1}] \text{ for } t \geq i \geq \tau) \quad (49)$$

By assumption $\lim_t \bar{S}_t = \infty$, so $\lim_t \bar{S}_t - \bar{S}_\tau = \infty$ as well. Clearly, the subtrahend in the display above is $o(\bar{S}_t - \bar{S}_\tau)$. Hence, taking the limit of $t \rightarrow \infty$, we have $\lim_t S_t = \infty$ with probability $1 - \delta$. Since δ was arbitrary, this also holds with probability one (by taking $\delta \rightarrow 0$). \square

Finally, we will need a classic result of Doob.

Lemma A.7 (Doob's Martingale Convergence Theorem (Doob, 2012)). *Given a random sequence $(X_t)_{t \geq 1}$ adapted to the filtration $(\mathcal{F}_t)_{t \geq 0}$, if $\forall t \geq 1 : \mathbb{E}[X_t | \mathcal{F}_{t-1}] \leq X_{t-1}$ and $\sup_{t \geq 0} \mathbb{E}[-\min(X_t, 0)] < \infty$, then $(X_t)_{t \geq 1}$ converges a.s. In particular, $X_t \rightarrow X$ a.s. as $t \rightarrow \infty$, where $X := \limsup_t X_t$ and $\mathbb{E}[|X|] < \infty$.*

B Bandits

In this section all results are stated in the bandit setting described in Section 3.1. We begin with a simple but crucial property of Algorithm 1, which follows from a symmetry of the update rule.

Lemma B.1 (Conservation of mass). *For all $t \geq 0$, $\sum_{a \in [K]} \theta_t(a) = \sum_{a \in [K]} \theta_0(a)$.*

Proof. Proceeding by induction, the base is tautological; recalling that a_t is the arm played at time t , we have

$$\sum_{a \in [K]} \theta_{t+1}(a) = \theta_{t+1}(a_t) + \sum_{a \in [K] \setminus \{a_t\}} \theta_{t+1}(a) \quad (50)$$

$$= \theta_t(a_t) + \eta(1 - \pi_t(a_t))r_t(a_t) + \sum_{a \in [K] \setminus \{a_t\}} [\theta_t(a) - \eta\pi_t(a)r_t(a_t)] \quad (51)$$

$$= \eta r_t(a_t) + \sum_{a \in [K]} [\theta_t(a) - \eta\pi_t(a)r_t(a_t)] \quad (52)$$

$$= \eta r_t(a_t) + \sum_{a \in [K]} \theta_t(a) - \eta r_t(a_t) \sum_{a \in [K]} \pi_t(a) \quad (53)$$

$$= \sum_{a \in [K]} \theta_t(a). \quad (\sum_{a \in [K]} \pi_t(a) = 1) \quad (54)$$

□

The rest of the proofs in this section will refer to the filtration $(\mathcal{F}_t)_{t \geq 0}$ defined by $\mathcal{F}_t := \sigma((a_i, r_i)_{i < t})$, and we adopt the shorthands $\mathbb{E}_t[\cdot] := \mathbb{E}[\cdot | \mathcal{F}_t]$ and $\text{Var}_t[\cdot] := \text{Var}[\cdot | \mathcal{F}_t]$. The following result is a stronger version of Lemma 2 of Mei et al. (2024a), and it guarantees that Algorithm 1 explores enough to keep trying all arms forever regardless of the observations.

Lemma 3.1 (Bandit Exploration). *Using Algorithm 1 with any constant learning rate $\eta \in \Theta(1)$, every arm is almost surely played infinitely often. That is, $\forall a \in [K] : N_\infty(a) = \infty$ almost surely.*

Proof. The first step is to show that, for any arm $b \in [K]$, if $|\{t \geq 0 : a_t = b\}| < \infty$ then $\sup_t |\theta_t(b)| < \infty$ a.s. Picking $b \in [K]$ and setting $m := \sup(\{0\} \cup \{t \geq 0 : a_t = b\})$, without assuming $|\{t \geq 0 : a_t = b\}| < \infty$ we have the bound

$$\sup_t |\theta_t(b)| \leq |\theta_0(b)| + \sup_t \sum_{i \in [t]} |\theta_i(b) - \theta_{i-1}(b)| \quad (\text{triangle inequality}) \quad (55)$$

$$\leq |\theta_0(b)| + \sum_{i \geq 1} |\theta_i(b) - \theta_{i-1}(b)| \quad (\sum_{i > t} |\theta_i(b) - \theta_{i-1}(b)| > 0) \quad (56)$$

$$= |\theta_0(b)| + \sum_{i \in [m]} |\theta_i(b) - \theta_{i-1}(b)| + \sum_{i > m} |\theta_i(b) - \theta_{i-1}(b)| \quad (57)$$

$$\leq |\theta_0(b)| + \sum_{i \in [m]} \eta R + \sum_{i > m} \eta R \pi_{\theta_i}(b) \quad (\text{update rule of Algorithm 1}) \quad (58)$$

$$\leq |\theta_0(b)| + \eta R(m + \sum_{t \geq 0} \pi_t(b)) \quad (\sum_{i=0}^m \eta R \pi_{\theta_i}(b) > 0) \quad (59)$$

$$=: \alpha(b). \quad (60)$$

Also, the Extended Borel-Cantelli Lemma (Lemma A.1) applied to $(\mathcal{F}_t)_{t \geq 0}$ with the event sequence $A_t := \{a_t = b\}$ implies

$$\sum_{t \geq 0} \mathbb{I}[a_t = b] < \infty \iff \sum_{t \geq 0} \pi_t(b) < \infty. \quad (61)$$

If $|\{t \geq 0 : a_t = b\}| < \infty$ then $m < \infty$ and $\sum_{t \geq 0} \mathbb{I}[a_t = b] < \infty$, and the latter inequality together with Eq. (61) implies $\sum_{t \geq 0} \pi_t(b) < \infty$ a.s., thus Eq. (60) yields $\sup_t |\theta_t(b)| \leq \alpha(b) < \infty$ a.s. A union bound over $b \in [K]$ implies that almost surely

$$\forall b \in [K] : |\{t \geq 0 : a_t = b\}| < \infty \implies \alpha(b) < \infty. \quad (62)$$

We are ready to fix an arm $a \in [K]$ and show that the event $\mathcal{E} := \{|\{t \geq 0 : a_t = a\}| < \infty\}$ has probability 0. For the remainder of the proof until the almost the very end we will work under the

assumption that \mathcal{E} occurs. On \mathcal{E} we have $\alpha(a) < \infty$ a.s, which implies $\sum_{t \geq 0} \pi_t(a) < \infty$ a.s, which in turn implies $\lim_t \pi_t(a) = 0$ a.s. The definition of $\pi_t(a)$ gives us

$$\lim_t \pi_t(a) = \lim_t \frac{\exp(\theta_t(a))}{\sum_{b \in [K]} \exp(\theta_t(b))} \quad (\text{Eq. (2)}) \quad (63)$$

$$\geq \lim_t \frac{\exp(-\alpha(a))}{\sum_{b \in [K]} \exp(\theta_t(b))} \quad (\alpha(a) \geq |\theta_t(a)|) \quad (64)$$

$$\geq \lim_t \frac{\exp(-\alpha(a))}{K \exp(\max_{b \in [K]} \theta_t(b))}, \quad (\sum_{b \in [K]} \exp(\theta_t(b)) \leq K \exp(\max_{b \in [K]} \theta_t(b))) \quad (65)$$

so from $\lim_t \pi_t(a) = 0$ a.s. we get $\lim_t \max_{b \in [K]} \theta_t(b) = \infty$ a.s. Then conservation of mass (Lemma B.1) implies that $\lim_t \min_{b \in [K]} \theta_t(b) = -\infty$ a.s. By Eq. (62) all arms that are selected only finitely often have parameters bounded away from $-\infty$ a.s, so there is a.s. an arm b that is played i.o. with $\liminf_t \theta_t(b) = -\infty$. We will refer to such an arm as b for the remainder of the proof. However, because b is played i.o, another application of the Extended Borel-Cantelli Lemma (to $(\mathcal{F}_t)_{t \geq 0}$ with events $A_t := \{a_t = b\}$) yields $\sum_{t \geq 0} \pi_t(b) = \infty$ a.s. Since $\sum_{t \geq 0} \pi_t(a) < \infty$ a.s, we have that $\pi_t(b) > \pi_t(a)$, and equivalently $\theta_t(b) > \theta_t(a)$, for infinitely many $t \geq 0$ a.s. In summary, $\theta_t(b)$ oscillates from being arbitrarily low to being larger than $\theta_t(a) \geq -\alpha(a)$.²

We will now argue that, for sufficiently large t , if $\theta_t(b) \leq \theta_t(a)$ but $\theta_{t+1}(b) > \theta_{t+1}(a)$ then $a_t = b$. Let T be the minimum timestep such that, for all $t \geq T$,

$$\max_{c \in [K]} \theta_t(c) \geq \log(\eta R) + \alpha(a), \quad \text{and} \quad a_t \neq a. \quad (66)$$

Since we are working on the event \mathcal{E} we have $a_t = a$ for only finitely many t , $\log(\eta R) + \alpha(a) < \infty$ a.s, and $\lim_t \max_{c \in [K]} \theta_t(c) = \infty$ a.s; taken together, these observations imply that $T < \infty$ exists a.s.

For $t \geq T$, suppose $\theta_t(b) \leq \theta_t(a)$ and $a_t \neq b$, and we must show

$$\theta_{t+1}(b) \leq \theta_{t+1}(a) \quad (67)$$

$$\iff \theta_t(b) - \eta \pi_t(b) r_t(a_t) \leq \theta_t(a) - \eta \pi_t(a) r_t(a_t) \quad (a_t \notin \{a, b\}) \quad (68)$$

$$\iff \eta r_t(a_t) (\pi_t(a) - \pi_t(b)) \leq \theta_t(a) - \theta_t(b). \quad (69)$$

Since $\theta_t(b) \leq \theta_t(a)$ we have $\pi_t(b) \leq \pi_t(a)$, and standard inequalities yield

$$\eta r_t(a_t) (\pi_t(a) - \pi_t(b)) \leq \eta R (\pi_t(a) - \pi_t(b)) \quad (0 \leq \pi_t(a) - \pi_t(b)) \quad (70)$$

$$= \eta R \frac{\exp(\theta_t(a)) - \exp(\theta_t(b))}{\sum_{c \in [K]} \exp(\theta_t(c))} \quad (71)$$

$$\leq \eta R \frac{\exp(\theta_t(a)) - \exp(\theta_t(b))}{\exp(\max_{c \in [K]} \theta_t(c))} \quad (\sum_{c \in [K]} \exp(\theta_t(c)) \geq \exp(\max_{c \in [K]} \theta_t(c))) \quad (72)$$

$$\leq \frac{\exp(\theta_t(a)) - \exp(\theta_t(b))}{\exp(\theta_t(a))} \quad (\text{Eq. (66), } \alpha(a) \geq \theta_t(a)) \quad (73)$$

$$= 1 - \exp(\theta_t(b) - \theta_t(a)) \quad (74)$$

$$\leq 1 - (1 + \theta_t(b) - \theta_t(a)). \quad (\exp(x) \geq 1 + x) \quad (75)$$

Thus Eq. (69) holds and we have established that, for all $t \geq T$, if $\theta_t(b) \leq \theta_t(a)$ and $\theta_{t+1}(b) > \theta_{t+1}(a)$ then $a_t = b$. Since $\theta_t(b)$ fluctuates from below $\theta_t(a)$ to above it i.o, we have that the events in the sequence $(B_t)_{t \geq 0}$ defined by $B_t := \{\theta_t(b) \leq \theta_t(a), a_t = b\} \in \mathcal{F}_{t+1}$ occur i.o. a.s. Applying the Extended Borel-Cantelli Lemma to $(\mathcal{F}_t)_{t \geq 0}$ and $(B_t)_{t \geq 0}$ implies that $\sum_{t \geq 0} \mathbb{P}(B_t | \mathcal{F}_t) = \infty$ a.s. However,

$$\mathbb{P}(B_t | \mathcal{F}_t) = \mathbb{I}[\theta_t(b) \leq \theta_t(a)] \pi_t(b) \leq \pi_t(a), \quad (76)$$

so $\sum_{t \geq 0} \mathbb{P}(B_t | \mathcal{F}_t) \leq \sum_{t \geq 0} \pi_t(a) < \infty$ a.s.

At this point we have that, on event \mathcal{E} , both $\sum_{t \geq 0} \mathbb{P}(B_t | \mathcal{F}_t) < \infty$ a.s. and $\sum_{t \geq 0} \mathbb{P}(B_t | \mathcal{F}_t) = \infty$ a.s. Since these events are mutually exclusive they both occur with probability 0, and since they are jointly exhaustive we have $\mathbb{P}(\mathcal{E}) = 0$. \square

²It is easy to see that $\theta_t(b)$ must also become arbitrarily large i.o, but this is not necessary for the proof.

The proof of our main result in the bandit setting, that $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \pi_t(a) = 1$, is broken into two propositions: the first guarantees that $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty$, in particular that as the time steps get large at least one $a \in \mathcal{A}^*$ will have an arbitrarily large parameter³; the second proposition says that $\lim_{t \rightarrow \infty} \theta_t(b) = -\infty$ for all $b \in [K] \setminus \mathcal{A}^*$. Taken together, the propositions imply that eventually some (potentially time step dependent) optimal arm dominates every suboptimal arm, establishing Convergence in Bandits (Theorem 3.2). We now turn to proving the two propositions.

The subsequent proofs will go a little smoother with some extra notation; we define $\Delta := \min_{a,b \in [K]: r(a) \neq r(b)} |r(a) - r(b)|$ to be the minimum nonzero gap between expected rewards of arms and $r(\mathcal{A}^*) := \max_{a \in [K]} r(a)$ to be the maximum attainable expected reward. Finally, we overload $\pi_t(\cdot)$ to take sets as input, i.e., given $\mathcal{S} \subset [K]$ we let $\pi_t(\mathcal{S}) := \sum_{a \in \mathcal{S}} \pi_t(a)$ be the probability that an arm in \mathcal{S} is selected. With these abbreviations in hand, the first proposition is as follows.

Proposition 3.3 (Infinite Optimal Parameters). *If $\mathcal{A}^* \neq [K]$ then $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty$ a.s.*

Proof. For $t \geq 0$, let $X_t := \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_t(a)$, such that $\sum_{i=0}^t X_i = \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_0(a)$. By the update rule of Algorithm 1, note also that

$$X_t = \eta \sum_{a \in \mathcal{A}^*} (\mathbb{I}[a_t = a] - \pi_t(a)) r_t. \quad (77)$$

The conditional expectation of X_t given \mathcal{F}_t can be lower bounded by

$$\mathbb{E}_t[X_t] = \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a] X_t] \quad (\sum_{a \in [K]} \mathbb{I}[a_t = a] = 1) \quad (78)$$

$$= \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a] \eta(1 - \pi_t(\mathcal{A}^*)) r_t] + \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b] \eta(-\pi_t(\mathcal{A}^*)) r_t] \quad (\text{Eq. (77)}) \quad (79)$$

$$= \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a] r_t] - \eta \pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b] r_t] \quad (\pi_t \text{ is } \mathcal{F}_t\text{-measurable}) \quad (80)$$

$$= \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \pi_t(a) r(a) - \eta \pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) r(b) \quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot] r_t] = \pi_t(\cdot) r(\cdot)) \quad (81)$$

$$\geq \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \pi_t(a) r(\mathcal{A}^*) - \eta \pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) (r(\mathcal{A}^*) - \Delta) \quad (r(a) = r(\mathcal{A}^*), r(b) \leq r(\mathcal{A}^*) - \Delta) \quad (82)$$

$$= \eta \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) (r(\mathcal{A}^*) - (r(\mathcal{A}^*) - \Delta)) \quad (\sum_{a \in \mathcal{A}^*} \pi_t(a) = \pi_t(\mathcal{A}^*), \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) = 1 - \pi_t(\mathcal{A}^*)) \quad (83)$$

$$= \eta \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) \Delta, \quad (84)$$

³Excluding the trivial case where $\mathcal{A}^* = [K]$, i.e. all arms are equally good.

and the conditional variance can be upper bounded by

$$\text{Var}_t[X_t] \leq \mathbb{E}_t[X_t^2] \quad (85)$$

$$= \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a] X_t^2] \quad (\sum_{a \in [K]} \mathbb{I}[a_t = a] = 1) \quad (86)$$

$$= \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a] (\eta(1 - \pi_t(\mathcal{A}^*)) r_t)^2] + \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b] (-\eta \pi_t(\mathcal{A}^*) r_t)^2] \quad (\text{Eq. (77)}) \quad (87)$$

$$\leq \eta^2 (1 - \pi_t(\mathcal{A}^*))^2 R^2 \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a]] + \eta^2 \pi_t(\mathcal{A}^*)^2 R^2 \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b]] \quad (r_t^2 \leq R^2) \quad (88)$$

$$= \eta^2 (1 - \pi_t(\mathcal{A}^*))^2 R^2 \sum_{a \in \mathcal{A}^*} \pi_t(a) + \eta^2 \pi_t(\mathcal{A}^*)^2 R^2 \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) \quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot]] = \pi_t(\cdot)) \quad (89)$$

$$= \eta^2 R^2 \left((1 - \pi_t(\mathcal{A}^*))^2 \pi_t(\mathcal{A}^*) + \pi_t(\mathcal{A}^*)^2 (1 - \pi_t(\mathcal{A}^*)) \right) \quad (\sum_{a \in \mathcal{A}^*} \pi_t(a) = \pi_t(\mathcal{A}^*), \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) = 1 - \pi_t(\mathcal{A}^*)) \quad (90)$$

$$= \eta^2 R^2 \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)). \quad (91)$$

Thus for all $t \geq 0$ we have $\text{Var}_t[X_t] \leq \eta R^2 \Delta^{-1} \mathbb{E}_t[X_t]$, $|X_t| \leq \eta R$, and X_t is \mathcal{F}_{t+1} -measurable. Setting $b := \eta R$, $\tau := 0$, and $c := \eta R^2 \Delta^{-1}$, we need only to prove that $\sum_{t \geq 0} \mathbb{E}_t[X_t] = \infty$, at which point we can apply the Freedman Divergence Trick (Lemma A.5) to conclude

$$\lim_t \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_0(a) = \sum_{t \geq 0} X_t = \infty \text{ a.s.} \quad (92)$$

$$\implies \lim_t \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty \text{ a.s.} \quad (\sum_{a \in \mathcal{A}^*} \theta_0(a) < \infty) \quad (93)$$

Thus in the remainder of the proof we turn our attention to showing $\sum_{t \geq 0} \mathbb{E}_t[X_t] = \infty$. Applying Eq. (84) and $\eta \Delta > 0$, we need only show that

$$\sum_{t \geq 0} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) = \infty. \quad (94)$$

Lemma 3.1 together with $\emptyset \neq \mathcal{A}^* \neq [K]$ implies that

$$\sum_{t \geq 0} \mathbb{I}[a_t \in \mathcal{A}^*] = \sum_{t \geq 0} \mathbb{I}[a_t \notin \mathcal{A}^*] = \infty \text{ a.s.} \quad (95)$$

Since $\mathbb{P}(a_t \in \mathcal{A}^* | \mathcal{F}_t) = \pi_t(\mathcal{A}^*)$ and $\mathbb{P}(a_t \notin \mathcal{A}^* | \mathcal{F}_t) = 1 - \pi_t(\mathcal{A}^*)$, the Extended Borel-Cantelli Lemma (Lemma A.1) applied to Eq. (95) furnishes $\sum_{t \geq 0} \pi_t(\mathcal{A}^*) = \sum_{t \geq 0} (1 - \pi_t(\mathcal{A}^*)) = \infty$ a.s. We now break into cases to show that Eq. (94) holds regardless of the behavior of $\pi_t(\mathcal{A}^*)$.

If $\pi_t(\mathcal{A}^*) \geq 1/2$ only finitely often then we can set $u := \max\{t \geq 0 : \pi_t(\mathcal{A}^*) \geq 1/2\}$ for

$$\sum_{t \geq 0} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) \geq \sum_{t > u} \frac{\pi_t(\mathcal{A}^*)}{2} = \infty. \quad (96)$$

Similarly, if $\pi_t(\mathcal{A}^*) < 1/2$ only finitely often then $u := \max\{t \geq 0 : \pi_t(\mathcal{A}^*) < 1/2\}$ gives us

$$\sum_{t \geq 0} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) \geq \sum_{t > u} \frac{1 - \pi_t(\mathcal{A}^*)}{2} = \infty. \quad (97)$$

We can narrow our focus to the case where $\pi_t(\mathcal{A}^*)$ is both above and below $1/2$ i.o. In particular, there must be infinitely many $t \geq 0$ such that $\pi_t(\mathcal{A}^*) < 1/2$ but $\pi_{\theta_{t+1}}(\mathcal{A}^*) \geq 1/2$, and for such t we have

$$\pi_{\theta_{t+1}}(\mathcal{A}^*) = \frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a))}{\sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a)) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_{t+1}(b))}. \quad (\text{Eq. (2)}) \quad (98)$$

The above equation is of the form $x/(x + y)$, where $x := \sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a))$ and $y := \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_{t+1}(b))$. Since $x/(x + y)$ is increasing in x and decreasing in y for $x, y > 0$, and $|\theta_{t+1}(c) - \theta_t(c)| \leq \eta R$ for all $c \in [K]$, we can maximize the right hand side for the upper bound

$$\pi_{\theta_{t+1}}(\mathcal{A}^*) \leq \frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R)}{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b) - \eta R)}. \quad (99)$$

Also, $\pi_t(\mathcal{A}^*) < 1/2$ yields $\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) < \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))$, so

$$\frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R)}{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b) - \eta R)} \quad (100)$$

$$= \frac{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))}{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) + \exp(-\eta R) \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))} \quad (101)$$

$$< \frac{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))}{(\exp(\eta R) + \exp(-\eta R)) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))} \quad (\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) < \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))) \quad (102)$$

$$= \frac{\exp(\eta R)}{\exp(\eta R) + \exp(-\eta R)} = \frac{\exp(2\eta R)}{\exp(2\eta R) + 1}. \quad (103)$$

Connecting the above displays, there are infinitely many $t \geq 0$ with $\pi_t(\mathcal{A}^*) < 1/2$ and $\pi_{\theta_{t+1}}(\mathcal{A}^*) \geq 1/2$, and for such t we have $1 - \pi_{\theta_{t+1}}(\mathcal{A}^*) > 1 - \exp(2\eta R)/(\exp(2\eta R) + 1) = (\exp(2\eta R) + 1)^{-1}$. Therefore $\pi_{\theta_{t+1}}(\mathcal{A}^*)(1 - \pi_{\theta_{t+1}}(\mathcal{A}^*)) \geq (2 \exp(2\eta R) + 2)^{-1}$ i.o, establishing Eq. (94). \square

The second proposition has a more complicated proof, due to the technical difficulty added by having multiple suboptimal arms with the same expected value. Controlling the suboptimal arms will be much more convenient with the following extra notation. Letting $n := |\{r(a) : a \in [K]\}|$ be the size of the range of the expected reward vector r , we partition the arms into $(\Phi_i)_{i \in [n]}$, where $\Phi_i := \arg \min_{a \in [K] \setminus \cup_{j < i} \Phi_j} r(a)$. Thus Φ_1 is the set of arms with minimal expected reward, Φ_2 is the set of arms with the second lowest expected reward, and so forth, culminating with $\Phi_n = \mathcal{A}^*$. Given $i \in [n]$, we will use the shorthands $\Phi_i^- := \cup_{j < i} \Phi_j$ and $\Phi_i^+ := \cup_{j > i} \Phi_j$. Note that Φ_i^- and Φ_i^+ are the sets of arms with lower, respectively higher, expected reward than the arms in Φ_i . First we will conjure up a couple bound that hold for the increments of suboptimal parameters.

Lemma B.2 (Bounds on the Expectation and Variance of Increments for Suboptimal Arms). *For any $i \in [n - 1]$, for any $b \in \Phi_i$, we have the bounds:*

$$\mathbb{E}_t[\theta_{t+1}(b) - \theta_t(b)] \leq \eta \pi_t(b) \left((1 - \pi_t(\Phi_i))r(b) - (r(b) + \Delta) \pi_t(\Phi_i^+) + R \pi_t(\Phi_i^-) \right), \quad (104)$$

$$\text{Var}_t[\theta_{t+1}(b) - \theta_t(b)] \leq \eta^2 R^2 \pi_t(b) (1 - \pi_t(b)). \quad (105)$$

Proof.

$$\begin{aligned} & \mathbb{E}_t[\theta_{t+1}(b) - \theta_t(b)] \\ &= \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a](\theta_{t+1}(b) - \theta_t(b))] \end{aligned} \tag{106}$$

$$\begin{aligned} &= \mathbb{E}_t[\mathbb{I}[a_t = b]\eta(1 - \pi_t(b))r_t] + \sum_{a \in [K] \setminus \{b\}} \mathbb{E}_t[\mathbb{I}[a_t = a]\eta(-\pi_t(b))r_t] \end{aligned} \tag{update rule of Algorithm 1} \tag{107}$$

$$\begin{aligned} &= \eta(1 - \pi_t(b))\mathbb{E}_t[\mathbb{I}[a_t = b]r_t] - \eta\pi_t(b) \sum_{a \in [K] \setminus \{b\}} \mathbb{E}_t[\mathbb{I}[a_t = a]r_t] \end{aligned} \tag{\pi_t is \mathcal{F}_t-measurable} \tag{108}$$

$$= \eta(1 - \pi_t(b))\pi_t(b)r(b) - \eta\pi_t(b) \sum_{a \in [K] \setminus \{b\}} \pi_t(a)r(a) \quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot]r_t] = \pi_t(\cdot)r(\cdot)) \tag{109}$$

$$\begin{aligned} &= \eta\pi_t(b) \left((1 - \pi_t(b))r(b) - \sum_{a \in \Phi_i \setminus \{b\}} \pi_t(a)r(b) - \sum_{c \in \Phi_i^+} \pi_t(c)r(c) - \sum_{d \in \Phi_i^-} \pi_t(d)r(d) \right) \end{aligned} \tag{110}$$

$$\begin{aligned} &\leq \eta\pi_t(b) \left((1 - \pi_t(b))r(b) - r(b) \sum_{a \in \Phi_i \setminus \{b\}} \pi_t(a) - (r(b) + \Delta) \sum_{c \in \Phi_i^+} \pi_t(c) + R \sum_{d \in \Phi_i^-} \pi_t(d) \right) \\ &\quad (r(c) \geq r(b) + \Delta, r(d) \geq -R) \tag{111} \end{aligned}$$

$$= \eta\pi_t(b) \left((1 - \pi_t(\Phi_i))r(b) - (r(b) + \Delta)\pi_t(\Phi_i^+) + R\pi_t(\Phi_i^-) \right). \tag{112}$$

$$\text{Var}_t[\theta_{t+1}(b) - \theta_t(b)] \leq \mathbb{E}_t[(\theta_{t+1}(b) - \theta_t(b))^2] \tag{113}$$

$$= \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a](\theta_{t+1}(b) - \theta_t(b))^2] \quad (\sum_{a \in [K]} \mathbb{I}[a_t = a] = 1) \tag{114}$$

$$\begin{aligned} &\leq \mathbb{E}_t[\mathbb{I}[a_t = b]\eta^2 R^2(1 - \pi_t(b))^2] + \sum_{a \in [K] \setminus \{b\}} \mathbb{E}_t[\mathbb{I}[a_t = a]\eta^2 R^2\pi_t(b)^2] \end{aligned} \tag{update rule of Algorithm 1} \tag{115}$$

$$\begin{aligned} &\leq \eta^2 R^2(1 - \pi_t(b))^2\pi_t(b) + \eta^2 R^2\pi_t(b)^2(1 - \pi_t(b)) \\ &\quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot]] = \pi_t(\cdot)) \tag{116} \end{aligned}$$

$$= \eta^2 R^2\pi_t(b)(1 - \pi_t(b)). \tag{117}$$

□

The next proposition will be applied inductively to control the relationship between the expectation and variance of arbitrary suboptimal arms.

Lemma B.3. *For constants $C, C' \geq 0$ and $i \in [n - 1]$, if all $b \in \Phi_i^-$ satisfy $\lim_{t \rightarrow \infty} \theta_t(b) = -\infty$ a.s. then there a.s. exists a finite timestep τ such that, for all $c \in \Phi_{i+1}^-$, $\sum_{a \in \mathcal{A}^*} \theta_\tau(a) \geq C + C'\theta_\tau(c)$.*

Proof. Without loss of generality suppose $C' \geq 1$. Throughout the proof we will use the following two constants, which depend on C and C' :

$$U_1 := C'(K\eta R + K \log(8RK/\Delta + K)) + C, \text{ and} \tag{118}$$

$$U_2 := \eta R + \log(8RK/\Delta + 1). \tag{119}$$

Fix $\epsilon \in (0, 1]$, and define

$$D := \sup_{x \geq 0} -x + 20\sqrt{1 + 4K^2\eta^2 R^2 + 4\eta R^2 x/\Delta} \log\left(\frac{2 + 4\eta R^2 x/\Delta}{\epsilon/K}\right), \tag{120}$$

noting that $D < \infty$. Proposition 3.3 says that $\lim_t \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty$ a.s, and we have by assumption that $\lim_t \theta_t(b) = -\infty$ a.s. for all $b \in \Phi_i^-$. Together these observations guarantee an a.s. finite

timestep μ such that

$$\sum_{a \in \mathcal{A}^*} \theta_\mu(a) \geq U_1 + D \geq -D \geq \max_{b \in \Phi_i^-} \theta_\mu(b). \quad (121)$$

Consider the collection of sequences $\{(X_t^b)_{t > \mu} : b \in \Phi_i^-\}$ defined by $X_t^b := \theta_t(b) - \theta_{t-1}(b)$, and the sequence $(Y_t)_{t > \mu}$ defined by $Y_t := \sum_{a \in \mathcal{A}^*} \theta_t(a) - \theta_{t-1}(a)$. For each $b \in \Phi_i^-$, $(X_t^b)_{t > \mu}$ satisfies the requirements of Freedman's Inequality (Lemma A.3) with $B := \eta R$; also, $(Y_t)_{t > \mu}$ does with $B := K\eta R$. Therefore we can apply Freedman's Inequality with $\delta := \epsilon/K$ to all of these sequences simultaneously and take a union bound to conclude that, with probability $1 - \epsilon(|\Phi_i^-| + 1)/K \leq 1 - \epsilon$,

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) - \theta_\mu(a) \geq \sum_{\mu < k \leq t} \mathbb{E}_{k-1}[Y_k^b] - 20 \sqrt{1 + 4K^2\eta^2R^2 + \sum_{\mu < k \leq t} \text{Var}_{k-1}[Y_k^b]} \log \left(\frac{2 + \sum_{\mu < k \leq t} \text{Var}_{k-1}[Y_k^b]}{\epsilon/K} \right), \quad (122)$$

$$\forall b \in \Phi_i^- : \theta_t(b) - \theta_\mu(b) \leq \sum_{\mu < k \leq t} \mathbb{E}_{k-1}[X_k^b] + 20 \sqrt{1 + 4\eta^2R^2 + \sum_{\mu < k \leq t} \text{Var}_{k-1}[X_k^b]} \log \left(\frac{2 + \sum_{\mu < k \leq t} \text{Var}_{k-1}[X_k^b]}{\epsilon/K} \right), \quad (123)$$

for all $t > \mu$. Let \mathcal{E} denote the event that both Eqs. (122) and (123) hold at all such t . We will argue that, on \mathcal{E} ,

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) \geq U_1 \geq 0 \geq \max_{b \in \Phi_i^-} \theta_t(b) \quad (124)$$

for all $t \geq \mu$ by strong induction. Thus let $t \geq \mu$, and suppose that Eq. (124) holds with k in place of t , for all $\mu \leq k < t$, noting that it holds for $k = \mu$ by the definition of μ . Eqs. (84) and (91) together imply that

$$\text{Var}_{k-1}[Y_k] \leq \frac{\eta R^2}{\Delta} \mathbb{E}_{k-1}[Y_t] \leq \frac{4\eta R^2}{\Delta} \mathbb{E}_{k-1}[Y_t], \quad (125)$$

for $k > \mu$, so the assumption that event \mathcal{E} holds implies

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) \geq \sum_{a \in \mathcal{A}^*} \theta_\mu(a) - D \quad (126)$$

$$\geq U_1 + D - D = U_1. \quad (\text{Eq. (121)}) \quad (127)$$

Now pick an arbitrary $b \in \Phi_i^-$. Without loss of generality, say $b \in \Phi_j$ for some $j < i$. For $\mu \leq k < t$ the inductive hypothesis implies that there exists $a \in \mathcal{A}^*$ such that $\theta_k(a) \geq U_1/K$. Thus

$$\pi_k(\Phi_j^+)/\pi_k(\Phi_j^-) \geq \pi_k(a)/\pi_k(\Phi_i^-) \quad (\Phi_j^- \subset \Phi_i^-, a \in \Phi_j^+) \quad (128)$$

$$\geq \exp(U_1/K)/K \quad (\max_{b \in \Phi_i^-} \theta_k(b) \leq 0) \quad (129)$$

$$\geq 1 + 4R/\Delta, \quad (130)$$

where the final inequality above follows from $U_1 \geq K \log(K + 4K/\Delta)$. Defining the constant $\gamma := (\Delta/2 + r(b) + R)/(\Delta + r(b) + R)$, we have $\gamma \leq (\Delta + 4R)/(2\Delta + 4R)$, which implies $\gamma/(1 - \gamma) \leq 1 + 4R/\Delta \leq \pi_k(\Phi_j^+)/\pi_k(\Phi_j^-)$. Therefore

$$\pi_k(\Phi_j^+) \geq \frac{\gamma}{1 - \gamma} \pi_k(\Phi_j^-) \quad (131)$$

$$\Rightarrow \pi_k(\Phi_j^+) \geq \gamma(\pi_k(\Phi_j^+) + \pi_k(\Phi_j^-)) \quad (132)$$

$$\Rightarrow \pi_k(\Phi_j^+)(\Delta + r(b) + R) \geq (\Delta/2 + r(b) + R)(1 - \pi_k(\Phi_j)) \quad (133)$$

$$\Rightarrow \pi_k(\Phi_j^+)(\Delta + r(b)) - R(1 - \pi_k(\Phi_j) - \pi_k(\Phi_j^+)) \geq (\Delta/2 + r(b))(1 - \pi_k(\Phi_j)) \quad (134)$$

$$\Rightarrow -\Delta(1 - \pi_k(\Phi_j))/2 \geq (1 - \pi_k(\Phi_j))r(b) - \pi_k(\Phi_j^+)(\Delta + r(b)) + R\pi_k(\Phi_j^-). \quad (135)$$

Combining Eqs. (104) and (135) produces

$$\mathbb{E}_{k-1}[X_k^b] \leq -\frac{\eta\Delta}{2} \pi_k(b)(1 - \pi_k(\Phi_j)). \quad (136)$$

From Eq. (130) we have $\pi_k(a) \geq \pi_k(\Phi_i^-) \geq \pi_k(\Phi_j)$, so $1 - \pi_k(\Phi_j) \geq 1/2$. Thus Eqs. (105) and (136) together provide

$$\mathbb{E}_{k-1}[X_k^b] \leq -\frac{\eta\Delta}{4}\pi_k(b) \quad (137)$$

$$\leq -\frac{\Delta}{4\eta R^2} \text{Var}_{k-1}[X_k^b]. \quad (138)$$

Eqs. (123) and (138) imply

$$\theta_t(b) = \theta_t(b) - \theta_\mu(b) + \theta_\mu(b) \quad (139)$$

$$\leq \theta_\mu(b) + D \quad (140)$$

$$\leq -D + D = 0. \quad (141)$$

Since $b \in \Phi_i^-$ was arbitrary, this concludes the inductive argument. We have shown that, on \mathcal{E} , Eq. (124) holds.

Define the stopping time ν by

$$\nu := \min \left\{ t \geq \mu : \left(\forall b \in \Phi_i : \theta_t(b) < U_2 \right) \text{ or } \left(\sum_{a \in \mathcal{A}^*} \theta_t(a) < U_1 \right) \text{ or } \left(\max_{b \in \Phi_i^-} \theta_t(b) > 0 \right) \right\}, \quad (142)$$

and define $(Z_t)_{t \geq \mu}$ by $Z_t := \sum_{b \in \Phi_i} \max(\theta_{\min(t, \nu)}(b), 0)$. We will show that $(Z_t)_{t \geq \mu}$ is a supermartingale, i.e. for all $t \geq \mu$, $\mathbb{E}_t[Z_{t+1} - Z_t] \leq 0$. If $t \geq \nu$ then we have $\mathbb{E}_t[Z_{t+1} - Z_t] = 0$, so assume $t < \nu$. Let $\mathcal{B} := \{b \in \Phi_i : \theta_t(b) \geq \eta R\}$ and $\mathcal{C} := \{c \in \Phi_i : \theta_t(c) < \eta R\}$, so

$$\mathbb{E}_t[Z_{t+1} - Z_t] = \sum_{b \in \Phi_i} \mathbb{E}_t[\max(\theta_{t+1}(b), 0) - \max(\theta_t(b), 0)] \quad (t < \nu) \quad (143)$$

$$= \sum_{b \in \mathcal{B}} \mathbb{E}_t[\max(\theta_{t+1}(b), 0) - \max(\theta_t(b), 0)] + \sum_{c \in \mathcal{C}} \mathbb{E}_t[\max(\theta_{t+1}(c), 0) - \max(\theta_t(c), 0)]. \quad (144)$$

The terms in the sum on the left of Eq. (144) can be bounded by

$$\mathbb{E}_t[\max(\theta_{t+1}(b), 0) - \max(\theta_t(b), 0)] = \mathbb{E}_t[\theta_{t+1}(b) - \theta_t(b)] \quad (145)$$

$$(|\theta_{t+1}(b) - \theta_t(b)| \leq \eta R, \theta_t(b) \geq 0 + \eta R)$$

$$\leq -\frac{\eta\Delta}{2}\pi_t(b)(1 - \pi_t(\Phi_i)), \quad (146)$$

where the last inequality above comes from Eq. (136) and the fact that $\nu > t$.⁴ For the sum on the right of Eq. (144), we can bound the terms by

$$\mathbb{E}_t[\theta_{t+1}(c) - \theta_t(c)] = \mathbb{E}_t[\mathbb{I}[a_t = c](\theta_{t+1}(c) - \theta_t(c))] + \mathbb{E}_t[\mathbb{I}[a_t \neq c](\theta_{t+1}(c) - \theta_t(c))] \quad (147)$$

$$\leq \mathbb{E}_t[\mathbb{I}[a_t = c]R\eta] + \mathbb{E}_t[\mathbb{I}[a_t \neq c]\pi_t(c)R\eta] \quad (\text{update rule of Algorithm 1}) \quad (148)$$

$$\leq 2\eta R\pi_t(c). \quad (149)$$

Combining Eqs. (144), (146) and (149) produces

$$\mathbb{E}_t[Z_{t+1} - Z_t] \leq -\frac{\eta\Delta}{2}(1 - \pi_t(\Phi_i)) \sum_{b \in \mathcal{B}} \pi_t(b) + 2\eta R \sum_{c \in \mathcal{C}} \pi_t(c) \quad (150)$$

$$= -\frac{\eta\Delta}{2}(1 - \pi_t(\Phi_i))\pi_t(\mathcal{B}) + 2\eta R\pi_t(\mathcal{C}). \quad (151)$$

⁴Specifically, $\nu > t$ implies the inductive hypothesis that was used to prove Eq. (136), and Φ_j can be replaced with Φ_i .

If $\pi_t(\mathcal{C}) = 0$ then the above is negative, so we may assume $\pi_t(\mathcal{C}) > 0$. Since $\nu > t$, there is some $b \in \Phi_i$ with $\theta_t(b) \geq U_2 \geq \eta R$, so

$$\pi_t(\mathcal{B})/\pi_t(\mathcal{C}) \geq \pi_t(b)/\pi_t(\mathcal{C}) \quad (152)$$

$$\geq \frac{\exp(\theta_t(b))}{\sum_{c \in \mathcal{C}} \exp(\theta_t(c))} \quad (153)$$

$$\geq \frac{\exp(U_2)}{n \exp(\eta R)} \quad (\text{definitions of } \mathcal{B} \text{ and } \mathcal{C}) \quad (154)$$

$$= \exp(U_2 - \eta R)/K \quad (155)$$

$$\geq (8RK/\Delta)/K \quad (U_2 \geq \eta R + \log(8RK/\Delta)) \quad (156)$$

$$= 8R/\Delta. \quad (157)$$

Also from $\nu > t$, we have that $\sum_{a \in \mathcal{A}^*} \theta_t(a) \geq U_1$, so at least one $a \in \mathcal{A}^*$ satisfies $\theta_t(a) \geq U_1/K$. Fixing such an a gives

$$(1 - \pi_t(\Phi_i))/\pi_t(\mathcal{C}) \geq \pi_t(a)/\pi_t(\mathcal{C}) \quad (158)$$

$$\geq \exp(U_1/K - \eta R)/K \quad (\text{like Eq. (155)}) \quad (159)$$

$$\geq (8Rn/\Delta + K)/K \quad (U_1 \geq K\eta R + K \log(8RK/\Delta + K)) \quad (160)$$

$$= 8R/\Delta + 1. \quad (161)$$

We will break into two cases, first assuming that $\pi_t(\Phi_i) \leq 1/2$. In this case we can upper bound Eq. (151) by

$$-\frac{\eta\Delta}{2}(1 - \pi_t(\Phi_i))\pi_t(\mathcal{B}) + 2\eta R\pi_t(\mathcal{C}) \leq -\frac{\eta\Delta}{4}\pi_t(\mathcal{B}) + 2\eta R\pi_t(\mathcal{C}) \quad (162)$$

$$\leq -2\eta R\pi_t(\mathcal{C}) + 2\eta R\pi_t(\mathcal{C}) \quad (\text{Eq. (157)}) \quad (163)$$

$$= 0. \quad (164)$$

On the other hand, if $\pi_t(\Phi_i) > 1/2$ then

$$1/2 < \pi_t(\mathcal{B}) + \pi_t(\mathcal{C}) \quad (165)$$

$$\leq \pi_t(\mathcal{B})(1 + \Delta/8R) \quad (\text{Eq. (157)}) \quad (166)$$

$$\Rightarrow \frac{2R}{4R + \Delta/2} \leq \pi_t(\mathcal{B}). \quad (167)$$

Starting once more from the right hand side of Eq. (151), we have

$$-\frac{\eta\Delta}{2}(1 - \pi_t(\Phi_i))\pi_t(\mathcal{B}) + 2\eta R\pi_t(\mathcal{C}) \leq -\frac{\Delta}{2} \cdot \frac{2\eta R}{4R + \Delta/2}(1 - \pi_t(\Phi_i)) + 2\eta R\pi_t(\mathcal{C}) \quad (\text{Eq. (167)}) \quad (168)$$

$$\leq -2\eta R \cdot \frac{\Delta/2}{4R + \Delta/2}(8R/\Delta + 1)\pi_t(\mathcal{C}) + 2\eta R\pi_t(\mathcal{C}) \quad (\text{Eq. (161)}) \quad (169)$$

$$= 0. \quad (170)$$

In concert, Eqs. (164) and (170) together with Eq. (151) imply that $\mathbb{E}_t[Z_{t+1} - Z_t] \leq 0$ when $\mu \leq t < \nu$. Therefore $(Z_t)_{t \geq \mu}$ is a submartingale, and it is clear from its definition that Z_t is bounded below by 0 at all times. We can apply Lemma A.7 and conclude that $(Z_t)_{t \geq \mu}$ converges a.s. to a random variable Z with $\mathbb{E}[|Z|] \leq \infty$.

We will again break into two cases, first assuming that $\nu = \infty$, i.e. the stopping time never stops. In this case $\lim_t Z_t = \lim_t \sum_{b \in \Phi_i} \max(\theta_t(b), 0)$, and this quantity will a.s. converge to a finite value; because each summand is nonnegative, this implies that all $b \in \Phi_i$ satisfy $\limsup_t \theta_t(b) < \infty$. From the assumption that $\forall c \in \Phi_i^- : \lim_t \theta_t(c) = -\infty$, we have that, for all $b \in \Phi_{i+1}^- = \Phi_i \cup \Phi_i^-$, $\limsup_t \theta_t(b) < \infty$. By Proposition 3.3, there a.s. exists a finite timestep τ such that $\sum_{a \in \mathcal{A}^*} \theta_\tau(a) \geq C + C' \max_{c \in \Phi_{i+1}^-} \limsup_t \theta_t(c) \geq C + C' \max_{c \in \Phi_{i+1}^-} \theta_\tau(c)$, as desired.

The other case is that $\nu < \infty$; this implies either the event \mathcal{E} fails to occur (since \mathcal{E} implies Eq. (124)), or $\forall b \in \Phi_i : \theta_\nu(b) < U_2$. On event \mathcal{E} , for all $b \in \Phi_{i+1}^-$,

$$C' \theta_\nu(b) + C \leq C' U_2 + C \leq U_1 \leq \sum_{a \in \mathcal{A}^*} \theta_\nu(a), \quad (0 + C, C' U_2 + C \leq U_1) \quad (171)$$

and setting $\tau := \nu$ gives the desired result. Therefore, regardless of whether or not $\nu < \infty$, the only way we don't have the desired result is if \mathcal{E} fails to occur, which happens with probability at most ϵ . Since ϵ was arbitrary, it can be taken to 0, and the desired result will hold a.s. \square

Having shown the above lemma, we are ready to establish that the parameters of suboptimal arms diverge to $-\infty$.

Proposition 3.4 (Negative Infinite Suboptimal Parameters). *For every suboptimal arm $b \in [K] \setminus \mathcal{A}^*$, $\lim_{t \rightarrow \infty} \theta_t(b) = -\infty$ a.s.*

Proof. Since $\Phi_n = \mathcal{A}^*$ and $\cup_{i \in [n]} \Phi_i = [K]$, the set of suboptimal arms is $\cup_{i \in [n-1]} \Phi_i$. Thus we will perform induction over $i \in [n-1]$, proving that all $b \in \Phi_i$ satisfy $\lim_t \theta_t(b) = -\infty$ a.s. from the inductive hypothesis that

$$\forall c \in \Phi_i^- : \lim_t \theta_t(c) = -\infty \text{ a.s.} \quad (172)$$

Note that Eq. (172) is vacuously satisfied for $i = 1$. Fix an arbitrary $\epsilon \in (0, 1]$, and define

$$D := \sup_{x \geq 0} -x + 20\sqrt{1 + 4K^2\eta^2R^2 + 4\eta R^2x/\Delta} \log\left(\frac{2 + 4\eta R^2x/\Delta}{\epsilon/K}\right), \quad (173)$$

noting that $D < \infty$. Let τ be the first timestep such that

$$\sum_{a \in \mathcal{A}^*} \theta_\tau(a) \geq (K+1)D + K \log(K+4K/\Delta) + K \max_{b \in \Phi_{i+1}^-} \theta_\tau(b), \quad (174)$$

and note that τ is a stopping time. Also, $\tau < \infty$ a.s. by applying Lemma B.3 (which is applicable due to the inductive hypothesis in Eq. (172)) with $C := (K+1)D + K \log(K+4K/\Delta)$ and $C' := K$.

now we can apply freedman's lemma to both the suboptimal arm and optimal sum, and conclude that the suboptimal arm goes to $-\infty$ wp $1 - \delta$. since δ was arbitrary the result becomes a.s. and the induction goes through meaning that the whole thing does.

Consider the collection of sequences $\{(X_t^b)_{t > \tau} : b \in \Phi_{i+1}^-\}$ defined by $X_t^b := \theta_t(b) - \theta_{t-1}(b)$, and the sequence $(Y_t)_{t > \tau}$ defined by $Y_t := \sum_{a \in \mathcal{A}^*} \theta_t(a) - \theta_{t-1}(a)$. For each $b \in \Phi_{i+1}^-$, $(X_t^b)_{t > \tau}$ satisfies the requirements of Freedman's Inequality (Lemma A.3) with $B := \eta R$; also, $(Y_t)_{t > \tau}$ does with $B := K\eta R$. Therefore we can apply Freedman's Inequality with $\delta := \epsilon/K$ to all of these sequences simultaneously and take a union bound to conclude that, with probability $1 - \epsilon(|\Phi_{i+1}^-| + 1)/K \leq 1 - \epsilon$,

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) - \theta_\tau(a) \geq \sum_{\tau < k \leq t} \mathbb{E}_{k-1}[Y_k^b] - 20 \sqrt{1 + 4K^2\eta^2R^2 + \sum_{\tau < k \leq t} \text{Var}_{k-1}[Y_k^b]} \log\left(\frac{2 + \sum_{\tau < k \leq t} \text{Var}_{k-1}[Y_k^b]}{\epsilon/K}\right), \quad (175)$$

$$\forall b \in \Phi_{i+1}^- : \theta_t(b) - \theta_\tau(b) \leq \sum_{\tau < k \leq t} \mathbb{E}_{k-1}[X_k^b] + 20 \sqrt{1 + 4\eta^2R^2 + \sum_{\tau < k \leq t} \text{Var}_{k-1}[X_k^b]} \log\left(\frac{2 + \sum_{\tau < k \leq t} \text{Var}_{k-1}[X_k^b]}{\epsilon/K}\right), \quad (176)$$

for all $t > \tau$. Let \mathcal{E} denote the event that both Eqs. (175) and (176) hold at all such t . We will argue that, on \mathcal{E} ,

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) \geq K \left(\log(K+4K/\Delta) + \max_{b \in \Phi_{i+1}^-} \theta_\tau(b) + D \right) \geq K \left(\log(K+4K/\Delta) + \max_{b \in \Phi_{i+1}^-} \theta_t(b) \right) \quad (177)$$

for all $t \geq \tau$ by strong induction. Thus let $t \geq \tau$, and suppose that Eq. (177) holds with k in place of t , for all $\tau \leq k < t$, noting that it holds for $k = \tau$ by the definition of τ . Eqs. (84) and (91) together imply that

$$\text{Var}_{k-1}[Y_k] \leq \frac{\eta R^2}{\Delta} \mathbb{E}_{k-1}[Y_k] \leq \frac{4\eta R^2}{\Delta} \mathbb{E}_{k-1}[Y_k], \quad (178)$$

for $k > \tau$, so the assumption that event \mathcal{E} holds implies

$$\sum_{a \in \mathcal{A}^*} \theta_t(a) \geq \sum_{a \in \mathcal{A}^*} \theta_\tau(a) - D \quad (179)$$

$$\geq K \left(\log(K + 4K/\Delta) + \max_{b \in \Phi_{i+1}^-} \theta_t(b) + D \right). \quad (180)$$

Now pick an arbitrary $b \in \Phi_{i+1}^-$. Without loss of generality, say $b \in \Phi_j$, where $j \in [i]$. For $\tau \leq k < t$ the inductive hypothesis implies that there exists $a \in \mathcal{A}^*$ such that $\theta_k(a) \geq \log(K + 4K/\Delta) + \max_{b \in \Phi_{i+1}^-} \theta_k(b)$. Thus

$$\pi_k(\Phi_j^+)/\pi_k(\Phi_j^-) \geq \pi_k(a)/\pi_k(\Phi_{i+1}^-) \quad (a \in \mathcal{A}^* \subset \Phi_{i+1}^+) \quad (181)$$

$$\geq \frac{\exp(\theta_k(a))}{K \exp(\max_{b \in \Phi_{i+1}^-} \theta_k(b))} \quad (182)$$

$$\geq \exp(\log(K + 4K/\Delta))/K \quad (183)$$

$$\geq 1 + 4R/\Delta. \quad (184)$$

Defining the constant $\gamma := (\Delta/2 + r(b) + R)/(\Delta + r(b) + R)$, we have $\gamma \leq (\Delta + 4R)/(2\Delta + 4R)$, which implies $\gamma/(1 - \gamma) \leq 1 + 4R/\Delta \leq \pi_k(\Phi_j^+)/\pi_k(\Phi_j^-)$. Therefore

$$\pi_k(\Phi_j^+) \geq \frac{\gamma}{1 - \gamma} \pi_k(\Phi_j^-) \quad (185)$$

$$\Rightarrow \pi_k(\Phi_j^+) \geq \gamma(\pi_k(\Phi_j^+) + \pi_k(\Phi_j^-)) \quad (186)$$

$$\Rightarrow \pi_k(\Phi_j^+)(\Delta + r(b) + R) \geq (\Delta/2 + r(b) + R)(1 - \pi_k(\Phi_j)) \quad (187)$$

$$\Rightarrow \pi_k(\Phi_j^+)(\Delta + r(b)) - R(1 - \pi_k(\Phi_j) - \pi_k(\Phi_j^+)) \geq (\Delta/2 + r(b))(1 - \pi_k(\Phi_j)) \quad (188)$$

$$\Rightarrow -\Delta(1 - \pi_k(\Phi_j))/2 \geq (1 - \pi_k(\Phi_j))r(b) - \pi_k(\Phi_j^+)(\Delta + r(b)) + R\pi_k(\Phi_j^-). \quad (189)$$

Combining Eqs. (104) and (189) produces

$$\mathbb{E}_{k-1}[X_k^b] \leq -\frac{\eta\Delta}{2} \pi_k(b)(1 - \pi_k(\Phi_j)). \quad (190)$$

From Eq. (184) we have $\pi_k(a) \geq \pi_k(\Phi_{i+1}^-) \geq \pi_k(\Phi_j)$, so $1 - \pi_k(\Phi_j) \geq 1/2$. Thus Eqs. (105) and (190) together provide

$$\mathbb{E}_{k-1}[X_k^b] \leq -\frac{\eta\Delta}{4} \pi_k(b) \quad (191)$$

$$\leq -\frac{\Delta}{4\eta R^2} \text{Var}_{k-1}[X_k^b]. \quad (192)$$

Eqs. (176) and (192) imply

$$\theta_t(b) = \theta_t(b) - \theta_\tau(b) + \theta_\tau(b) \quad (193)$$

$$\leq D + \max_{b \in \Phi_{i+1}^-} \theta_\tau(b), \quad (194)$$

and multiplying both sides of Eq. (194) by K before adding $K \log(K + 4K/\Delta)$ implies the second inequality of Eq. (177) (since $b \in \Phi_{i+1}^-$ was arbitrary). This concludes the inductive argument over $t \geq \tau$. We have shown that, on \mathcal{E} , Eq. (177) holds. In fact, on event \mathcal{E} , we can also use Eqs. (176) and (192) together with the fact that $\sum_{t \geq 1} \text{Var}_{t-1}[X_t^b] = \infty$ (using Eq. (105)) to conclude that $\lim_t \theta_t(b) = -\infty$ for an arbitrary $b \in \Phi_{i+1}^-$. This finishes off the inductive argument over $i \in [n-1]$. \square

The above results are all that is needed for the proof of Theorem 3.2. Next, we show that the stochastic gradient bandit algorithm (Algorithm 1) only converges to “generalized one-hot policies”, i.e. $\sum_{a \in \mathcal{A}^*} \pi_t(a) = 1$, and not true “one-hot policies”, i.e. $\exists a \in \mathcal{A}^* : \pi_t(a) = 1$. Among the optimal arms, there will be permanent non-stationary behavior.

Theorem B.4 (Theorem 1.4, Bramson et al. (2004)). *Suppose $(X_t)_{t \geq 0}$ is a sub-martingale with increments $(I_t)_{t \geq 0}$, satisfying*

$$\mathbb{E}_{t-1}[I_t^- \mathbb{I}\{I_t^- > x\}] \geq \mathbb{E}_{t-1}[I_t^+ \mathbb{I}\{I_t^+ > bx\}] \quad (195)$$

almost surely, for all $t \geq 1$ and fixed $x \geq x_1$, for a fixed b and $x_1 > 0$. Then,

$$\mathbb{P}(\text{either } (X_t)_{t \geq 0} \text{ converges or } \limsup_{t \rightarrow \infty} X_t = \infty) = 1 \quad (196)$$

Lemma B.5 (Finite total quadratic variations). *Let $\{X_t\}_{t \geq 0}$ be a discrete-time sub-martingale and $I_t := X_t - X_{t-1}$ for all $t \geq 1$ to be an increment at time t . If $\{X_t\}_{t \geq 0}$ converges a.s and $|I_t| < \infty$ for all t , then $\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[I_s^2] < \infty$.*

Proof. Let $\{X_t\}_{t \geq 0}$ be a discrete-time sub-martingale. Define an increment at time t by $I_t := X_t - X_{t-1}$ for all $t \geq 1$. Note that $X_t = \sum_{s=1}^t I_s$. Also, $|I_t| < \infty$ for all t . Suppose $\{X_t\}_{t \geq 0}$ converges a.s. By Doob decomposition, X_t can be uniquely written as

$$X_t = M_t + A_t \quad (197)$$

where $A_t := \sum_{s=1}^t \mathbb{E}_{s-1}[I_s]$ is predictable and non-decreasing, and $M_t := \sum_{s=1}^t (I_s - \mathbb{E}_{s-1}[I_s])$ is a martingale. Let $Y_t := I_t - \mathbb{E}_{t-1}[I_t]$ be an increment at time t . Note that Y_s is bounded, i.e $|Y_s| < \infty$ for all s . Since $X_t = M_t + A_t$ converges a.s to a finite limit, both M_t and A_t must remain finite a.s. In particular, A_t cannot diverge to ∞ because that would force M_t to diverge to $-\infty$. However, the increment Y_t are bounded for all t . Hence M_t cannot diverge to $-\infty$. Therefore, $\{A_t\}_{t \geq 0}$ and $\{M_t\}_{t \geq 0}$ converge a.s. Since A_t converges and is a non-decreasing, then $\mathbb{E}_{t-1}[I_t]$ must converge to 0 a.s. Specifically, there must exist a timestep τ such that for all $s \geq \tau$, we have $0 \leq \mathbb{E}_{s-1}[I_s] < 1$. In other words, we have $\mathbb{E}_{s-1}[I_s]^2 \leq \mathbb{E}_{s-1}[I_s]$ for all $s \geq \tau$. Since $\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[I_s] < \infty$, we know that $\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[I_s]^2 < \infty$. Since $\{M_t\}_{t \geq 0}$ converges a.s, then M_t is pathwise bounded a.s

$$\Pr(\sup_t |M_t| < \infty) = 1 \quad (198)$$

Let define a stopping time $\tau_K := \inf\{t \geq 1 : |M_t| \geq K\}$ and a corresponding stopped martingale $Z_t = M_{t \wedge \tau_K}$. For all $t < \tau_K$, $|Z_t| = |M_t| < K$. When $t = \tau_K$, then $|Z_t| = |M_{\tau_K}| = |M_{\tau_K-1} + Y_{\tau_K}| \leq |M_{\tau_K-1}| + |Y_{\tau_K}| < \infty$ since Y_t is bounded for all t . Therefore, Z_t is a uniformly bounded martingale, i.e there exists a constant C , s.t $|Z_t| \leq C$ a.s

$$\exists C \text{ s.t } P(\sup_t |Z_t| \leq C) = 1 \quad (199)$$

Denote $Y'_t := Z_t - Z_{t-1}$ as an increment of the martingale Z_t , i.e $Z_t = \sum_{s=1}^t Y'_s$. Note that $Y'_t = Y_t \mathbb{1}\{t \leq \tau_K\}$. Since Z_t is a bounded martingale, then Z_t is also bounded in L^2 . Specifically,

$$\sup_t \mathbb{E}[Z_t^2] \leq \sup_t \mathbb{E}[C^2] = C^2 \quad \text{a.s} \quad (200)$$

Note that

$$\mathbb{E}[Z_{t+1}^2 - Z_t^2] = \mathbb{E}[(Z_t + Y'_{t+1})^2 - Z_t^2] \quad (201)$$

$$= \mathbb{E}[2Y'_{t+1}Z_t] + \mathbb{E}[Y'^2_{t+1}] \quad (202)$$

$$= 2\mathbb{E}[Z_t \mathbb{E}_t[Y'^2_{t+1}]] + \mathbb{E}[Y'^2_{t+1}] \quad \text{Law of total expectation} \quad (203)$$

$$= \mathbb{E}[Y'^2_{t+1}] \quad (204)$$

Recursively, we have $\mathbb{E}[Z_t^2] - \mathbb{E}[Z_0^2] = \sum_{s=1}^t \mathbb{E}[Y'^2_s] = \mathbb{E}[\sum_{s=1}^t \mathbb{E}_{s-1}[Y'^2_s]]$. Since $\sup_t \mathbb{E}[Z_t^2] \leq C'^2$ a.s, then

$$\mathbb{E}[\sum_{s=1}^t \mathbb{E}_{s-1}[Y'^2_s]] \text{ exists and is finite a.s} \quad (205)$$

Since $\sum_{s=1}^t \mathbb{E}_{s-1}[Y'^2_s]$ is non-negative and finite on expectation, then

$$\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[Y'^2_s] < \infty \quad \text{a.s} \quad (206)$$

Note that $Y'_s = Y_s 1\{s \leq \tau_K\}$. Then,

$$\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[(Y_s 1\{s \leq \tau\})^2] = \sum_{s=1}^{\infty} 1\{s \leq \tau_K\} \mathbb{E}_{s-1}[Y_s^2] = \sum_{s=1}^{\infty \wedge \tau_K} \mathbb{E}_{s-1}[Y_s^2] < \infty \quad (207)$$

In other words, for all K , the total predictable variation of martingale M_t is finite. Therefore, $\sum_{s=1}^{\infty} \mathbb{E}_{s-1}[Y_s^2] < \infty$ a.s. Therefore, $\sum_{s=1}^{\infty} \text{Var}_{s-1}[I_s] = \sum_{s=1}^{\infty} \mathbb{E}_{s-1}[I_s^2] < \infty$. \square

Lemma B.6 (Infinite total variance). *In the bandit setting, where we allow ties in the expected reward of arms, using Algorithm 1 with any $\eta \in \Omega(1)$, we have that for all $a \in \mathcal{A}^*$, almost surely,*

$$\sum_{t=0}^{\infty} \pi_t(a)(1 - \pi_t(a)) = \infty \quad (208)$$

Proof. We divide into three cases: $\pi_t(a)$ is finitely often above $\frac{1}{2}$, finitely often below $\frac{1}{2}$, and infinitely often above and below $\frac{1}{2}$. First, by Lemma 3.1, we know that for all $a \in [K]$, $\sum_{t \geq 0} \pi_t(a) = \sum_{t \geq 0} (1 - \pi_t(a)) = \infty$. In the first case, since $\pi_t(a)$ is only finitely often above $\frac{1}{2}$,

$$\sum_{t \geq 0} \pi_t(a)(1 - \pi_t(a)) \geq \sum_{t \geq 0} \mathbb{I}\{\pi_t(a) \leq \frac{1}{2}\} \frac{\pi_t(a)}{2} = \infty \quad (209)$$

The inequality is due to $x(1-x) \geq \frac{x}{2}$ when $x \leq \frac{1}{2}$. Similarly, when $\pi_t(a)$ is finitely often below $\frac{1}{2}$,

$$\sum_{t \geq 0} \pi_t(a)(1 - \pi_t(a)) \geq \sum_{t \geq 0} \mathbb{I}\{\pi_t(a) > \frac{1}{2}\} \frac{1 - \pi_t(a)}{2} = \infty \quad (210)$$

The inequality is due to $x(1-x) > \frac{1-x}{2}$ when $x > \frac{1}{2}$. To show the last case, it is equivalently to show that the event $\pi_t(a) < \frac{1}{2}$ and $\pi_{\theta_{t+1}}(a) > \frac{1}{2}$ happens i.o. Since $\pi_t(a) = \frac{\exp(\theta_t(a))}{\exp(\theta_t(a)) + \sum_{a' \neq a} \exp(\theta_t(a'))}$, we denote $X := \exp(\theta_t(a))$ and $Y := \sum_{a' \neq a} \exp(\theta_t(a'))$. Since $\pi_t(a) < \frac{1}{2}$, we have $X < Y$. In order to increase $\pi_{\theta_{t+1}}(a) > \frac{1}{2}$, the algorithm needs to play $a_t = a$, so the update rule is

$$\theta_{t+1}(a) = \theta_t(a) + \eta(1 - \pi_t(a))r_t(a) \leq \theta_t(a) + \eta R \quad (211)$$

For other actions $a' \neq a$, the update rule is

$$\theta_{t+1}(a') = \theta_t(a') - \eta\pi_t(a')r_t(a) \geq \theta_t(a') - \eta R \quad (212)$$

Since $\pi_t(a) = \frac{X}{X+Y}$ is an increasing function in X when $X, Y > 0$, then

$$\pi_{\theta_{t+1}}(a) = \frac{\exp(\theta_{t+1}(a))}{\exp(\theta_{t+1}(a)) + \sum_{a' \neq a} \exp(\theta_{t+1}(a'))} \quad (213)$$

$$\leq \frac{\exp(\theta_t(a) + \eta R)}{\exp(\theta_t(a) + \eta R) + \sum_{a' \neq a} \exp(\theta_t(a'))} \quad (214)$$

$$\leq \frac{\exp(\theta_t(a) + \eta R)}{\exp(\theta_t(a) + \eta R) + \sum_{a' \neq a} \exp(\theta_t(a') - \eta R)} \quad (215)$$

$$= \frac{X \exp(\eta R)}{X \exp(\eta R) + Y \exp(-\eta R)} \quad (216)$$

$$\leq \frac{\exp(\eta R)}{\exp(\eta R) + \exp(-\eta R)} \quad (X < Y)$$

Therefore, $\pi_{\theta_{t+1}}(a)(1 - \pi_{\theta_{t+1}}(a)) \geq \frac{\exp(-\eta R)}{\exp(\eta R) + \exp(-\eta R)}$. Hence, $\sum_{t \geq 0} \pi_{\theta_{t+1}}(a)(1 - \pi_{\theta_{t+1}}(a)) = \infty$. \square

Proposition 2.2 (Non-Stationary Convergence). *In the bandit setting, where the mean reward has ties, using Algorithm 1 with any $\eta \in \Theta(1)$, for all $a \in \mathcal{A}^*$,*

$$\limsup_t \theta_t(a) = \infty \text{ a.s.} \quad (1)$$

In other words, $(\pi_t)_{t \geq 0}$ does not converge to any one-hot policy.

Proof. Denote the incremental $I_t(a) := \eta(\mathbb{I}\{a_t = a\} - \pi_t(a))r_t(a_t)$. By the update rule (Algorithm 1), we know that

$$\theta_{t+1}(a) = \theta_t(a) + \eta(\mathbb{I}\{a_t = a\} - \pi_t(a))r_t(a_t) = \theta_t(a) + I_t(a) \quad (217)$$

Also, we have,

$$\mathbb{E}_t[I_t(a)] = \eta\mathbb{E}_t[(\mathbb{I}\{a_t = a\} - \pi_{\theta_t}(a))r_t(a_t)] \quad (218)$$

$$= \eta\pi_{\theta_t}(a)(1 - \pi_{\theta_t}(a))r(a) - \sum_{a' \neq a} \eta\pi_{\theta_t}(a')\pi_{\theta_t}(a)r(a') \quad (219)$$

$$= \pi_{\theta_t}(a)(r(a) - \pi_{\theta_t}^\top r) \quad (220)$$

$$\geq 0 \quad (221)$$

Therefore, $\mathbb{E}_t[\theta_{t+1}(a)] \geq \theta_t(a)$. Hence, $(\theta_t(a))_{t \geq 0}$ is a sub-martingale. Note that $I_t(a) \in [-\eta R, \eta R]$ for all t . Therefore, by setting $x = \eta R > 0$ and $b = 2 > 0$, then $\mathbb{E}\{I_t^+ > bx\} = \mathbb{E}\{I_t^- > x\} = 0$. Hence, $\mathbb{E}_{t-1}[I_t^-\mathbb{I}\{I_t^- > x\}] = \mathbb{E}_{t-1}[I_t^+\mathbb{I}\{I_t^+ > bx\}] = 0$. In other words, the condition of Theorem B.4 is satisfied trivially. Therefore, $\{\theta_t(a)\}_{t \geq 0}$ can either converges to a finite value or $\limsup_{t \rightarrow \infty} \theta_t(a) = \infty$ a.s. Suppose $\{\theta_t(a)\}_{t \geq 0}$ converges a.s. By Lemma B.5, $\sum_{t=0}^{\infty} \mathbb{E}_{t-1}[I_t^2] < \infty$. In other words, $\sum_{t=0}^{\infty} \pi_t(a)(1 - \pi_t(a)) < \infty$. However, by Lemma B.6, we know that $\sum_{t=0}^{\infty} \pi_t(a)(1 - \pi_t(a)) = \infty$. Therefore, $\limsup_t \theta_t(a) = \infty$ a.s. \square

C Nonstationary Bandit Setting

Proposition C.1 (Infinite Optimal Parameters). *If $\mathcal{A}^* \neq [K]$ then $\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty$ almost surely.*

Proof. For $t \geq 0$, let $X_t := \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_t(a)$, such that $\sum_{i=0}^t X_i = \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_0(a)$. By the update rule of Algorithm 1, note also that

$$X_t = \eta \sum_{a \in \mathcal{A}^*} (\mathbb{I}[a_t = a] - \pi_t(a))r_t. \quad (222)$$

The conditional expectation of X_t given \mathcal{F}_t can be lower bounded by

$$\mathbb{E}_t[X_t] = \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a]X_t] \quad (\sum_{a \in [K]} \mathbb{I}[a_t = a] = 1) \quad (223)$$

$$= \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a]\eta(1 - \pi_t(\mathcal{A}^*))r_t] + \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b]\eta(-\pi_t(\mathcal{A}^*))r_t] \quad (\text{Eq. (222)}) \quad (224)$$

$$= \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a]r_t] - \eta\pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b]r_t] \quad (\pi_t \text{ is } \mathcal{F}_t\text{-measurable}) \quad (225)$$

$$= \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \pi_t(a)r^t(a) - \eta\pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b)r^t(b) \quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot]r_t] = \pi_t(\cdot)r^t(\cdot)) \quad (226)$$

$$\geq \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \pi_t(a)(r(a) - \Delta/3) - \eta\pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b)(r(b) + \Delta/3) \quad (\forall a \in [K], \forall t \geq T, |r^t(a) - r(a)| \leq \Delta/3) \quad (227)$$

$$\geq \eta(1 - \pi_t(\mathcal{A}^*)) \sum_{a \in \mathcal{A}^*} \pi_t(a)(r(\mathcal{A}^*) - \Delta/3) - \eta\pi_t(\mathcal{A}^*) \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b)(r(\mathcal{A}^*) - 2\Delta/3) \quad (r(a) = r(\mathcal{A}^*), r(b) \leq r(\mathcal{A}^*) - \Delta) \quad (228)$$

$$= \eta\pi_t(\mathcal{A}^*)(1 - \pi_t(\mathcal{A}^*))(r(\mathcal{A}^*) - \Delta/3 - (r(\mathcal{A}^*) - 2\Delta/3)) \quad (\sum_{a \in \mathcal{A}^*} \pi_t(a) = \pi_t(\mathcal{A}^*), \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) = 1 - \pi_t(\mathcal{A}^*)) \quad (229)$$

$$= \eta\pi_t(\mathcal{A}^*)(1 - \pi_t(\mathcal{A}^*))\Delta/3, \quad (230)$$

and the conditional variance can be upper bounded by

$$\text{Var}_t[X_t] \leq \mathbb{E}_t[X_t^2] \quad (231)$$

$$= \sum_{a \in [K]} \mathbb{E}_t[\mathbb{I}[a_t = a] X_t^2] \quad (\sum_{a \in [K]} \mathbb{I}[a_t = a] = 1) \quad (232)$$

$$= \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a] (\eta(1 - \pi_t(\mathcal{A}^*)) r_t)^2] + \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b] (-\eta \pi_t(\mathcal{A}^*) r_t)^2] \quad (\text{Eq. (77)}) \quad (233)$$

$$\leq \eta^2 (1 - \pi_t(\mathcal{A}^*))^2 R^2 \sum_{a \in \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = a]] + \eta^2 \pi_t(\mathcal{A}^*)^2 R^2 \sum_{b \in [K] \setminus \mathcal{A}^*} \mathbb{E}_t[\mathbb{I}[a_t = b]] \quad (r_t^2 \leq R^2) \quad (234)$$

$$= \eta^2 (1 - \pi_t(\mathcal{A}^*))^2 R^2 \sum_{a \in \mathcal{A}^*} \pi_t(a) + \eta^2 \pi_t(\mathcal{A}^*)^2 R^2 \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) \quad (\mathbb{E}_t[\mathbb{I}[a_t = \cdot]] = \pi_t(\cdot)) \quad (235)$$

$$= \eta^2 R^2 \left((1 - \pi_t(\mathcal{A}^*))^2 \pi_t(\mathcal{A}^*) + \pi_t(\mathcal{A}^*)^2 (1 - \pi_t(\mathcal{A}^*)) \right) \quad (\sum_{a \in \mathcal{A}^*} \pi_t(a) = \pi_t(\mathcal{A}^*), \sum_{b \in [K] \setminus \mathcal{A}^*} \pi_t(b) = 1 - \pi_t(\mathcal{A}^*)) \quad (236)$$

$$= \eta^2 R^2 \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)). \quad (237)$$

Thus for all $t \geq \tau$ we have $\text{Var}_t[X_t] \leq \eta^2 R^2 \Delta^{-1} \mathbb{E}_t[X_t]$, $|X_t| \leq \eta R$, and X_t is \mathcal{F}_{t+1} -measurable. Setting $b := \eta R$ and $c := \eta R^2 \Delta^{-1}$, we need only to prove that $\sum_{t \geq \tau} \mathbb{E}_t[X_t] = \infty$, at which point we can apply the Freedman Divergence Trick (Lemma A.5) to conclude

$$\lim_t \sum_{a \in \mathcal{A}^*} \theta_{t+1}(a) - \theta_0(a) = \sum_{t \geq \tau} X_t = \infty \text{ a.s.} \quad (238)$$

$$\implies \lim_t \sum_{a \in \mathcal{A}^*} \theta_t(a) = \infty \text{ a.s.} \quad (\sum_{a \in \mathcal{A}^*} \theta_0(a) < \infty) \quad (239)$$

Thus in the remainder of the proof we turn our attention to showing $\sum_{t \geq \tau} \mathbb{E}_t[X_t] = \infty$. Applying Eq. (230) and $\eta \Delta / 3 > 0$, we need only show that

$$\sum_{t \geq \tau} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) = \infty. \quad (240)$$

Lemma 3.1 together with $\emptyset \neq \mathcal{A}^* \neq [K]$ implies that

$$\sum_{t \geq \tau} \mathbb{I}[a_t \in \mathcal{A}^*] = \sum_{t \geq \tau} \mathbb{I}[a_t \notin \mathcal{A}^*] = \infty \text{ a.s.} \quad (241)$$

Since $\mathbb{P}(a_t \in \mathcal{A}^* | \mathcal{F}_t) = \pi_t(\mathcal{A}^*)$ and $\mathbb{P}(a_t \notin \mathcal{A}^* | \mathcal{F}_t) = 1 - \pi_t(\mathcal{A}^*)$, the Extended Borel-Cantelli Lemma (Lemma A.1) applied to Eq. (95) furnishes $\sum_{t \geq \tau} \pi_t(\mathcal{A}^*) = \sum_{t \geq \tau} (1 - \pi_t(\mathcal{A}^*)) = \infty$ a.s. We now break into cases to show that Eq. (240) holds regardless of the behavior of $\pi_t(\mathcal{A}^*)$.

If $\pi_t(\mathcal{A}^*) \geq 1/2$ only finitely often then we can set $u := \max\{t \geq 0 : \pi_t(\mathcal{A}^*) \geq 1/2\}$ for

$$\sum_{t \geq \tau} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) \geq \sum_{t > u} \frac{\pi_t(\mathcal{A}^*)}{2} = \infty. \quad (242)$$

Similarly, if $\pi_t(\mathcal{A}^*) < 1/2$ only finitely often then $u := \max\{t \geq 0 : \pi_t(\mathcal{A}^*) < 1/2\}$ gives us

$$\sum_{t \geq \tau} \pi_t(\mathcal{A}^*) (1 - \pi_t(\mathcal{A}^*)) \geq \sum_{t > u} \frac{1 - \pi_t(\mathcal{A}^*)}{2} = \infty. \quad (243)$$

We can narrow our focus to the case where $\pi_t(\mathcal{A}^*)$ is both above and below $1/2$ i.o. In particular, there must be infinitely many $t \geq 0$ such that $\pi_t(\mathcal{A}^*) < 1/2$ but $\pi_{\theta_{t+1}}(\mathcal{A}^*) \geq 1/2$, and for such t we have

$$\pi_{\theta_{t+1}}(\mathcal{A}^*) = \frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a))}{\sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a)) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_{t+1}(b))}. \quad (\text{Eq. (2)}) \quad (244)$$

The above equation is of the form $x/(x + y)$, where $x := \sum_{a \in \mathcal{A}^*} \exp(\theta_{t+1}(a))$ and $y := \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_{t+1}(b))$. Since $x/(x + y)$ is increasing in x and decreasing in y for $x, y > 0$, and $|\theta_{t+1}(c) - \theta_t(c)| \leq \eta R$ for all $c \in [K]$, we can maximize the right hand side for the upper bound

$$\pi_{\theta_{t+1}}(\mathcal{A}^*) \leq \frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R)}{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b) - \eta R)}. \quad (245)$$

Also, $\pi_t(\mathcal{A}^*) < 1/2$ yields $\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) < \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))$, so

$$\frac{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R)}{\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a) + \eta R) + \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b) - \eta R)} \quad (246)$$

$$= \frac{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))}{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) + \exp(-\eta R) \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))} \quad (247)$$

$$< \frac{\exp(\eta R) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))}{(\exp(\eta R) + \exp(-\eta R)) \sum_{a \in \mathcal{A}^*} \exp(\theta_t(a))} \quad (\sum_{a \in \mathcal{A}^*} \exp(\theta_t(a)) < \sum_{b \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(b))) \quad (248)$$

$$= \frac{\exp(\eta R)}{\exp(\eta R) + \exp(-\eta R)} = \frac{\exp(2\eta R)}{\exp(2\eta R) + 1}. \quad (249)$$

Connecting the above displays, there are infinitely many $t \geq 0$ with $\pi_t(\mathcal{A}^*) < 1/2$ and $\pi_{\theta_{t+1}}(\mathcal{A}^*) \geq 1/2$, and for such t we have $1 - \pi_{\theta_{t+1}}(\mathcal{A}^*) > 1 - \exp(2\eta R)/(\exp(2\eta R) + 1) = (\exp(2\eta R) + 1)^{-1}$. Therefore $\pi_{\theta_{t+1}}(\mathcal{A}^*)(1 - \pi_{\theta_{t+1}}(\mathcal{A}^*)) \geq (2\exp(2\eta R) + 2)^{-1}$ i.o, establishing Eq. (240). \square

Proposition C.2 (Finite Suboptimal Parameters). *For every suboptimal arm $b \in [K] \setminus \mathcal{A}^*$, $\lim_{t \rightarrow \infty} \theta_t(b) = -\infty$ a.s.*

Remark C.3. The proof remains virtually unchanged from the proof of Proposition 3.4, and the necessary changes are identical to the ones made for the proof of Proposition C.1.

Theorem C.4. *In the non-stationary bandit setting described as above, Algorithm 1 with any $\eta \in \Theta(1)$ almost surely converges to playing optimal arms,*

$$\lim_{t \rightarrow \infty} \sum_{a \in \mathcal{A}^*} \pi_t(a) \rightarrow 1 \text{ a.s.} \quad (250)$$

D Reinforcement Learning

Define the MDP $\mathcal{M} = (\mathcal{H}, \mathcal{S}, \mathcal{A}, \{r_h\}_{h=0}^{H-1}, \{P_h\}_{h=0}^{H-1}, \rho)$. Let $N_t(s, a) := \sum_{t \geq 0} \mathbb{I}\{s_t = s, a_t = a\}$ be the total number of visitations of state-action pair (s, a) until episode t . We denote that $\mathbb{P}_t^{h+1}(s_{h+1} = s' | s_h = s)$ as the probability of visiting state s' in the horizon $h+1$ from the state s in the horizon h during the episode t . First, we extend the bandit exploration lemma (Lemma 3.1) to obtain its counterpart in the RL setting.

Lemma D.1 (RL exploration (Lemma 4.1)). *Using the REINFORCE algorithm with any $\eta \in \Theta(1)$ under the finite-horizon MDP \mathcal{M} defined as above, for all $h \in \mathcal{H}$, for all reachable $s \in \mathcal{S}_h$ and for all $a \in \mathcal{A}_s$ we have, almost surely, that every reachable state action pair will be visited i.o, i.e $N_\infty(s, a) = \infty$.*

Proof. First, for all $h \in \mathcal{H}$, for a given reachable $s \in \mathcal{S}_h$ that is played infinite often, every action $a \in \mathcal{A}_s$ will be played i.o. by the bandit exploration result (Lemma 3.1). In other words, for all $h \in \mathcal{H}$, for a reachable state s that is played i.o, we have, almost surely that,

$$N_\infty(s, a) = \infty \iff \sum_{t \geq 0} \pi_t^h(a|s) = \infty \quad \forall a \in \mathcal{A}_s \quad (251)$$

Next, for all $h \in \mathcal{H}$, we want to show that every reachable state $s \in \mathcal{S}_h$ will be visited i.o. by induction. Suppose for a given $h \in \mathcal{H}$, for some reachable $s \in \mathcal{S}_h$ and there exists an action $a \in \mathcal{A}_s$

such that $P_{h+1}(s_{h+1} = s' | s_h = s, a_h = a) > 0$ for some $s' \in \mathcal{S}_{h+1}$, if s is visited i.o, s' is also visited i.o. For the base case $h = 0$, for some reachable states $s \in \mathcal{S}_0$, i.e $\rho(s) > 0$ we have,

$$\sum_{t \geq 0} \rho(s) = \infty \quad (252)$$

since $\rho(s)$ is a constant for every episode. Therefore, every reachable state $s \in \mathcal{S}_0$ is visited i.o. For the inductive case, if any reachable states $s \in \mathcal{S}_h$ is visited i.o, then any reachable states $s' \in \mathcal{S}_{h+1}$ is also visited i.o. A state s' is reachable if there exists an action $a \in \mathcal{A}_s$ from a reachable state $s \in \mathcal{S}_h$ such that $P_{h+1}(s'|s, a) > 0$. Denote $c := \min_{s \in \mathcal{S}_h} \min_{a \in \mathcal{A}_s} P_{h+1}(s'|s, a)$ be the minimum transition probability from the horizon h to $h+1$ among states and actions. For reachable s' from s , we have .

$$\sum_{t \geq 0} \mathbb{P}_t^{h+1}(s'|s) = \sum_{t \geq 0} \sum_{a \in \mathcal{A}_s} P_{h+1}(s'|s, a) \pi_t^h(a|s) \quad (253)$$

$$\geq \sum_{a \in \mathcal{A}_s} c \sum_{t \geq 0} \pi_t^h(a|s) \\ = \infty \quad (\text{by Eq. (251)}) \quad (254)$$

Therefore, if $s \in \mathcal{S}_h$ is reachable and visited i.o, then any reachable states $s' \in \mathcal{S}_{h+1}$ from s will be visited i.o. Combined Eq. (251) and Eq. (253), for all $h \in \mathcal{H}$, we have that any reachable state-action $(s, a) \in \mathcal{S}_h \times \mathcal{A}_h$ pairs will be visited i.o. we know that every state-action pair will be visited i.o. \square

Next, we obtain the convergence of REINFORCE in the finite-horizon setting.

Theorem D.2 (RL convergence (Theorem 4.2)). *For the MDP defined as above, using the algorithm REINFORCE with constant learning rate $\eta \in \Theta(1)$, we have, almost surely, for all $s \in \mathcal{S}_0$, $V_0^{\pi_t}(s) \rightarrow V_0^*(s)$ as $t \rightarrow \infty$.*

Proof. We denote $\delta := \min_s \min_{a,b \in \mathcal{A}_s, a \neq b} |Q(s, a) - Q(s, b)| > 0$ to be the minimum non-zero gap between Q -values. Denote $\mathcal{A}_h^* = \{a | a = \arg \max_{a \in \mathcal{A}_s} r(s, a)\}$ is the set of optimal action at a given state s . We also denote $C := \max_s \max_a \min_b (Q(s, a) - Q(s, b))$. We want to prove by backward induction that for all reachable state $s_0 \in \mathcal{S}_0$, $\sum_{a \in \mathcal{A}_0^*} \pi_t^0(a|s) \rightarrow 1$ as $t \rightarrow \infty$. Suppose for all $h' \in \{h, \dots, H-1\}$, for all reachable $s \in \mathcal{S}_{h'}$, we have $\sum_{a \in \mathcal{A}_{h'}^*} \pi_t^{h'}(a|s) \rightarrow 1$ as $t \rightarrow \infty$, we want to prove that for all reachable $s \in \mathcal{S}_{h-1}$, we have $\sum_{a \in \mathcal{A}_{h-1}^*} \pi_t^{h-1}(a|s) \rightarrow 1$ as $t \rightarrow \infty$. In the base case $h = H-1$, the REINFORCE update rule (Algorithm 2) is reduced to,

$$\theta_{t+1}^{H-1}(s, a) = \theta_t^{H-1}(s, a) + \eta(\mathbb{I}[a_{H-1} = a] - \pi_t^{H-1}(a|s))r_h \quad (255)$$

This is the bandit update rule (Algorithm 1) for a given reachable state $s \in \mathcal{S}_{H-1}$. By Theorem 3.2, for a given reachable state $s \in \mathcal{S}_{H-1}$, using the stochastic gradient bandit algorithm with constant learning rate $\eta \in \Theta(1)$, we will have, almost surely that $\sum_{a^* \in \mathcal{A}_h^*} \pi_t^{H-1}(a^*|s) \rightarrow 1$ as $t \rightarrow \infty$. By Lemma D.1 , any reachable states $s \in \mathcal{S}_{H-1}$ will be sampled i.o. Hence, using REINFORCE with $\eta \in \Theta(1)$, that for all reachable $s \in \mathcal{S}_{H-1}$ that are played i.o, we have, almost surely, $\sum_{a \in \mathcal{A}_{H-1}^*} \pi_t^{H-1}(a|s) \rightarrow 1$ as $t \rightarrow \infty$. In other words, $V_{H-1}^{\pi_t}(s) \rightarrow V_{H-1}^*(s)$ as $t \rightarrow \infty$.

For inductive case, suppose for all $h' \in \{h, \dots, H-1\}$, for all reachable $s \in \mathcal{S}_{h'}$, we have $\sum_{a \in \mathcal{A}_{h'}^*} \pi_t^{h'}(a|s) \rightarrow 1$ as $t \rightarrow \infty$, we want to prove that for all reachable $s \in \mathcal{S}_{h-1}$, we have $\sum_{a \in \mathcal{A}_{h-1}^*} \pi_t^{h-1}(a|s) \rightarrow 1$ as $t \rightarrow \infty$. By the induction hypothesis, for all $h \in \mathcal{H}$, for all reachable $s \in \mathcal{S}_h$, $V_h^{\pi_t}(s) \rightarrow V_h^*(s)$ and $Q_h^{\pi_t}(s, a) \rightarrow Q_h^*(s, a)$ for all $a \in \mathcal{A}_s$ as $t \rightarrow \infty$. First, we note that

$$V_{h'}^*(s) - V_{h'}^{\pi_t}(s) = \sum_{a'} \pi_t^{h'}(a'|s) (\max_a Q_{h'}^*(s, a) - Q_{h'}^{\pi_t}(s, a')) \quad (256)$$

$$= \sum_{a'} \pi_t^{h'}(a'|s) \underbrace{(\max_a Q_{h'}^*(s, a) - Q_{h'}^*(s, a'))}_{C_1} + \underbrace{Q_{h'}^*(s, a') - Q_{h'}^{\pi_t}(s, a')}_{C_2} \quad (257)$$

We denote that $\eta_{h'}(t) := \sum_{a' \notin \mathcal{A}_h^*} \pi_t^{h'}(a'|s)$. For the first term C_1 , we have

$$C_1 = \sum_{a'} \pi_t^{h'}(a'|s) (\max_{a \in \mathcal{A}_s} Q_{h'}^*(s, a) - Q_{h'}^*(s, a')) \quad (258)$$

$$= \sum_{a' \notin \mathcal{A}_h^*} \pi_t^{h'}(a'|s) (\max_{a \in \mathcal{A}_s} Q_{h'}^*(s, a) - Q_{h'}^*(s, a')) \quad (259)$$

$$\leq C\gamma_{h'}(t) \quad (260)$$

since the horizon H is fixed and $r_h \leq R$ for all $h \in \mathcal{H}$, then

$$\max_{a \in \mathcal{A}_s} Q_{h'}^*(s, a) - Q_{h'}^*(s, a') \leq C \quad (261)$$

By the induction hypothesis, we have, for all $h' \in \{h, \dots, H-1\}$, we have that $\gamma_{h'}(t) \rightarrow 0$ as $t \rightarrow \infty$.

For the second term C_2 , we have $Q_{h'}^*(s, a') - Q_{h'}^{\pi_t}(s, a') \leq \alpha_{h'}(t)$, where $\alpha_{h'}(t) \rightarrow 0$ as $t \rightarrow \infty$ by induction hypothesis. Therefore,

$$V_{h'}^*(s) - V_{h'}^{\pi_t}(s) \leq C\gamma_{h'}(t) + \alpha_{h'}(t) \quad (262)$$

Denote $\epsilon_h(t) := C\gamma_{h'}(t) + \alpha_{h'}(t)$ and $\epsilon_h(t) \rightarrow 0$ as $t \rightarrow \infty$. Hence, for sufficiently large timestep τ , such that for all $t \geq \tau$, for all $h' = h, \dots, H-1$, for all reachable $s \in \mathcal{S}_{h'}$, we have that

$$V_{h'}^*(s) - V_{h'}^{\pi_t}(s) \leq \frac{\delta}{3} \quad (263)$$

where δ is the minimum possible gap Q-value defined as above. The existence of τ is guaranteed since $\epsilon_h(t) \rightarrow 0$ as $t \rightarrow \infty$. First, for a given reachable $s \in \mathcal{S}_{h-1}$, such that for any actions $a \in \mathcal{A}_s$, we have,

$$Q_{h-1}^{\pi_t}(s, a) = r_{h-1}(s, a) + \mathbb{E}_{s' \sim P_h(\cdot|s, a)}[V_h^{\pi_t}(s')] \quad (264)$$

$$\geq r_{h-1}(s, a) + \mathbb{E}_{s' \sim P_h(\cdot|s, a)}[V_h^*(s')] - \epsilon \quad (265)$$

$$= Q_{h-1}^*(s, a) - \epsilon \quad (266)$$

Also, for any actions $a \in \mathcal{A}_s$, we have,

$$Q_{h-1}^{\pi_t}(s, a) \leq Q_{h-1}^*(s, a) \leq Q_{h-1}^*(s, a) + \epsilon \quad (267)$$

By definition of δ , for a given reachable $s \in \mathcal{S}_{h-1}$, we know that $Q_{h-1}^*(s, a) - Q_{h-1}^*(s, b) \geq \delta$ for any $a, b \in \mathcal{A}_s$ such that $a \neq b$, then we have

$$Q_{h-1}^{\pi_t}(s, a) - Q_{h-1}^{\pi_t}(s, b) \geq \delta - 2\epsilon \geq \frac{\delta}{3} \quad (268)$$

Note that δ is a non-zero gap by definition. Note that the update rule of Algorithm 2,

$$\theta_{t+1}^{h-1}(s, a) = \theta_{h-1}^{h-1}(s, a) + \eta(1\{a_{h-1} = a\} - \pi_t^{h-1}(a|s)) \sum_{h'=h-1}^{H-1} r_{h'} \quad (269)$$

is equivalent to the update rule in the nonstationary bandit setting, by considering only the updates to the arms at state s and the full trajectory's rewards as the observed rewards.

Since by definition

$$\mathbb{E}^{\pi_t} \left[\sum_{h'=h-1}^{H-1} r_{h'} \middle| s_{h-1} = s, a_{h-1} = a \right] = Q_{h-1}^{\pi_t}(s, a),$$

$\sum_{h'=h-1}^{H-1} r_{h'}$ is an unbiased estimate of $Q_{h-1}^{\pi_t}(s_{h-1}, a_{h-1})$, up to nonstationarity in π_t which eventually diminishes below $\delta/3$. Also, note that

$$Q_{h-1}^{\pi_t}(s, a) = \mathbb{E}^{\pi_t} \left[\sum_{h'=h-1}^{H-1} r_{h'} \middle| s_{h-1} = s, a_{h-1} = a \right] \leq R(H-h), \quad (270)$$

since $r(s, a) \leq R, \forall s \in \mathcal{S}, a \in \mathcal{A}$. Since the sample return $\sum_{h'=h-1}^{H-1} r_{h'}$ is a bounded, unbiased estimator of $Q_{h-1}^{\pi^t}(s_{h-1}, a_{h-1})$, and there is a minimum gap of $\delta/3$ between Q values among different actions within the same state, we can apply the convergence result from the nonstationary bandit setting (Theorem C.4) to conclude that $\sum_{a \in \mathcal{A}_{h-1}^*(s)} \pi_t^{h-1}(a|s) \rightarrow 1$ as $t \rightarrow \infty$ for all reachable $s \in \mathcal{S}_{h-1}$. Therefore the induction hypothesis holds, and we conclude that, using REINFORCE with $\eta \in \Theta(1)$, $\sum_{a \in \mathcal{A}_h} \pi_t^h(a|s) \rightarrow 1$ as $t \rightarrow \infty$ for all $s \in \mathcal{S}_h$ (or $V_0^{\pi^t}(s) \rightarrow V_0^*(s)$ as $t \rightarrow \infty$ for all $s \in \mathcal{S}_0$). \square

E Convergence rate

To obtain the convergence rate of the REINFORCE algorithm (Algorithm 2), we first generalize the convergence rate result from the bandit setting with the uniqueness assumption to the one without it. Then, we also obtain the convergence rate in the non-stationary bandit setting before showing the rate of the REINFORCE algorithm.

Theorem E.1. *In the bandit setting where multiple arms can have a same reward, for a large enough τ , for all $T > \tau$, the average sub-optimality decreases at a rate $O(\frac{\log T}{T})$. Formally, for a constant c , we have*

$$\frac{1}{T} \sum_{s=\tau}^T \left(r(a^*) - \langle \pi_{\theta_s}, r \rangle \right) \leq \frac{c \log(T - \tau)}{(T - \tau)} \quad (271)$$

Proof. By Eq. (84), we have

$$\mathbb{E}_t[\theta_{t+1}(\mathcal{A}^*) - \theta_t(\mathcal{A}^*)] \geq \eta \Delta \pi_{\theta_t}(\mathcal{A}^*)(1 - \pi_{\theta_t}(\mathcal{A}^*)) \geq 0 \quad (272)$$

By Theorem 3.2, we have $\lim_t \pi_{\theta_t}(\mathcal{A}^*) = 1$ a.s. Therefore, for a large enough t , we have

$$\pi_{\theta_t}(\mathcal{A}^*) \geq \frac{1}{2} \quad (273)$$

By Lemma 3.1, we know that every action $a \in [K]$ will be played i.o. In other words, for all $a \in [K]$, $\sum_{t \geq 0} \pi_t(a) = \infty$. Therefore, we have

$$\sum_{t=0}^{\infty} (1 - \pi_{\theta_t}(\mathcal{A}^*)) = \infty \quad (274)$$

Therefore, we have

$$\sum_{t=0}^{\infty} \mathbb{E}_t[\theta_{t+1}(\mathcal{A}^*) - \theta_t(\mathcal{A}^*)] = \infty \quad \text{a.s} \quad (275)$$

By Eq. (91), we have

$$\text{Var}_t[\theta_{t+1}(\mathcal{A}^*) - \theta_t(\mathcal{A}^*)] \leq \eta^2 R^2 \pi_t(\mathcal{A}^*)(1 - \pi_t(\mathcal{A}^*)) \quad (276)$$

Since the conditional expectation and variance of the bound sequence $\{\theta_{t+1}(\mathcal{A}^*) - \theta_t(\mathcal{A}^*)\}_{t \geq 0}$ are proportional, we can use the Lemma A.5 to show that the expectation will dominate the variance eventually. Therefore, for all large enough $t \geq \tau$, for some constant $C > 0$

$$\frac{1}{|\mathcal{A}^*|} \theta_t(\mathcal{A}^*) \geq C \sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*)) \quad (277)$$

It is easy to see that $\sup_t \theta_t(a) < \infty$ for all $a \in [K] \setminus \mathcal{A}^*$. Therefore, for a large enough $t \geq \tau$, we have

$$\theta_t(a) - \frac{1}{|\mathcal{A}^*|} \theta_t(\mathcal{A}^*) \leq -C \sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*)) \quad (278)$$

which implies that

$$\sum_{a \in [K] \setminus \mathcal{A}^*} \exp(\theta_t(a) - \frac{1}{|\mathcal{A}^*|} \theta_t(\mathcal{A}^*)) \leq (K - |\mathcal{A}^*|) \exp(-C \sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*))) \quad (279)$$

Therefore, we have

$$1 - \pi_{\theta_t}(\mathcal{A}^*) \leq \frac{1 - \pi_{\theta_t}(\mathcal{A}^*)}{\pi_{\theta_t}(\mathcal{A}^*)} \quad (280)$$

$$= \sum_{a \in [K] \setminus \mathcal{A}^*} \frac{\pi_{\theta_t}(a)}{\pi_{\theta_t}(\mathcal{A}^*)} \quad (281)$$

$$= \sum_{a \in [K] \setminus \mathcal{A}^*} \frac{\exp(\theta_t(a))}{\sum_{a' \in \mathcal{A}^*} \exp(\theta_t(a'))} \quad (282)$$

$$\leq \sum_{a \in [K] \setminus \mathcal{A}^*} \frac{\exp(\theta_t(a))}{|\mathcal{A}^*| \exp(\frac{1}{|\mathcal{A}^*|} \theta_t(\mathcal{A}^*))} \quad (\text{Jensen's inequality}) \quad (283)$$

$$\leq (K - |\mathcal{A}^*|) \exp(-C \sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*))) \quad (\text{Eq. (279)}) \quad (284)$$

By (Mei et al., 2024a, Lemma 15) with $x_n = \sum_{s=\tau}^{t-1} (1 - \pi_{\theta_s}(\mathcal{A}^*)) > 0$, $x_{n+1} = \sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*)) > 0$, $c = C > 0$, $B = (K - |\mathcal{A}^*|) \geq 1$, gives us for all $t \geq \tau$,

$$\sum_{s=\tau}^t (1 - \pi_{\theta_s}(\mathcal{A}^*)) \leq \frac{1}{C} \log(C(t - \tau) + \exp(CM)) + \frac{\pi^2}{12C} \quad (285)$$

where $M = \max\{B, \frac{1}{C} \log(C(K - |\mathcal{A}^*|)), 1 - \pi_{\theta_\tau}(\mathcal{A}^*)\}$. Finally, for all $s \geq \tau$ and $T > \tau$, we have

$$r(a^*) - \langle \pi_{\theta_s}, r \rangle = \sum_{a \in [K] \setminus \mathcal{A}^*} \pi_{\theta_s}(a)(r(a^*) - r(a)) \leq 2R(1 - \pi_{\theta_s}(\mathcal{A}^*)) \quad (286)$$

Summing from τ to T , we have

$$\frac{1}{T} \sum_{s=\tau}^T \left(r(a^*) - \langle \pi_{\theta_s}, r \rangle \right) \leq \frac{2R(\frac{1}{C} \log(C(T - \tau) + \exp(CM)) + \frac{\pi^2}{12C})}{T - \tau} \quad (287)$$

□

Theorem E.2. *In the non-stationary bandit setting, for a large enough τ , then for all $T > \tau$, the average sub-optimality decreases at a rate $O(\frac{\log T}{T})$. Formally, for a constant c , we have*

$$\frac{1}{T} \sum_{s=\tau}^T \left(r(a^*) - \langle \pi_{\theta_s}, r \rangle \right) \leq \frac{c \log(T - \tau)}{T - \tau} \quad (288)$$

Proof. Repeating the same analysis with Eq. (230), Eq. (237), Theorem C.4, we have, for all $t \geq \tau''$,

$$\sum_{s=\tau''}^t (1 - \pi_{\theta_s}(\mathcal{A}^*)) \leq \frac{1}{C} \log(C(t - \tau) + \exp(CM)) + \frac{\pi^2}{12C} \quad (289)$$

where $M = \max\{B, \frac{1}{C} \log(C(K - |\mathcal{A}^*|)), 1 - \pi_{\theta_{\tau''}}(\mathcal{A}^*)\}$. Also, from the non-stationary bandit setting, there exists τ' such that for all $t \geq \tau'$,

$$|r(a) - r^t(a)| \leq \frac{\Delta}{3} \quad (290)$$

for all $a \in [K]$. Therefore, for all $s \geq \max\{\tau', \tau''\}$ and $T > \max\{\tau', \tau''\}$, we have

$$r(a^*) - \langle \pi_{\theta_s}, r \rangle \leq 2R(1 - \pi_{\theta_s}(\mathcal{A}^*)) \quad (291)$$

Summing from $\tau := \max\{\tau', \tau''\}$ to T , we have

$$\frac{1}{T} \sum_{s=\tau}^T \left(r(a^*) - \langle \pi_{\theta_s}, r \rangle \right) \leq \frac{2R(\frac{1}{C} \log(C(T - \tau) + \exp(CM)) + \frac{\pi^2}{12C})}{T - \tau} \quad (292)$$

□

Theorem E.3. *In the finite-horizon MDP setting, for a large enough τ , for all $T > \tau$, for all $s \in \mathcal{S}_0$, the average sub-optimality decreases at a rate $O(\frac{\log T}{T})$. Formally, for a constant c , we have*

$$\frac{1}{T} \sum_{s=\tau}^T \left(V_0^*(s) - V_0^{\pi_s}(s) \right) \leq \frac{c \log T}{T} \quad (293)$$

Proof. Repeating the same analysis, we have, for each $h \in \{0, \dots, H-1\}$, for all $s \in \mathcal{S}_h$, for all $t \geq \tau_h$,

$$\sum_{s=\tau_h}^t (1 - \pi_h^s(\mathcal{A}_s^*|s)) \leq \frac{1}{C} \log(C(t - \tau_h) + \exp(CM)) + \frac{\pi^2}{12C} \quad (294)$$

where $M_h = \max\{|\mathcal{A}_s| - |\mathcal{A}_s^*|, \frac{1}{C} \log(C(|\mathcal{A}_s| - |\mathcal{A}_s^*|)), 1 - \pi_{\theta_{\tau_h}}(\mathcal{A}_s^*|s)\}$. Also, there exists τ'_h such that for all $t \geq \tau'_h$,

$$\|Q_h^*(s, \cdot) - Q_h^{\pi_t}(s, \cdot)\|_\infty \leq \frac{\Delta}{3} \quad (295)$$

for all $a \in \mathcal{A}_s$. Therefore, for all horizon $h \in \{0, \dots, H-1\}$, for all $t \geq \max\{\tau_h, \tau'_h\}$ and $T > \max\{\tau_h, \tau'_h\}$, we have

$$Q_h^*(s, a^*) - \sum_{a \in \mathcal{A}_s} \pi_h^t(a|s) Q_h^{\pi_t}(s, a) \leq Q_h^*(s, a^*) - \sum_{a \in \mathcal{A}_s} \pi_h^t(a|s) (Q_h^*(s, a) - \frac{\Delta}{3}) \quad (296)$$

$$= \sum_{a \in \mathcal{A}_s \setminus \mathcal{A}_s^*} \pi_h^t(a|s) (Q_h^*(s, a^*) - Q_h^*(s, a)) + \frac{\Delta}{3} \quad (297)$$

$$\leq 2(H-h)R(1 - \pi_h^t(\mathcal{A}_s^*|s)) + \frac{\Delta}{3} \quad (298)$$

which implies

$$\frac{1}{T} \sum_{s=\max\{\tau_h, \tau'_h\}}^T (V_h^*(s) - V_h^{\pi_t}(s)) \leq \frac{2R(H-h)(\frac{1}{C} \log(C(T - \tau_h) + \exp(CM)) + \frac{\pi^2}{12C}) + \frac{\Delta}{3}(T - \tau_h)}{T - \tau_h} \quad (299)$$

Since for all $h \in \{0, \dots, H-1\}$, for all $s \in \mathcal{S}_h$, for all $a \in \mathcal{A}_s$, $\lim_t Q_h^{\pi_t}(s, a) = Q_h^*(s, a)$, we can take $\delta \rightarrow 0$. Therefore,

$$\frac{1}{T} \sum_{s=\max\{\tau_0, \tau'_0\}}^T (V_0^*(s) - V_0^{\pi_t}(s)) \leq \frac{2HR\frac{1}{C} \log(C(T - \tau_0) + \exp(CM)) + \frac{\pi^2}{12C}}{T - \tau_0} \quad (300)$$

□

F Additional experiments

Specifically, we measured the average suboptimality in the last episodes over 30 of the algorithm in longer ChainMDP, DeepSea environment and CartPole environment. For ChainMDP, we extended the lengths of the environment to $H = \{4, 5, 6\}$ and measured the average suboptimality gap across 100 learning rates (from $\exp(-9)$ to $\exp(1)$). For each length, we observed a clear bowl-shaped curve. As complexity (chain length) increased, the specific thresholds of the bowl shape varied slightly, but the optimal learning rate remained consistently around $\eta \approx .95$. Next, we gradually increase the complexity of our evaluation by testing the REINFORCE algorithm (Algorithm 2) on the deep sea treasure environment. The agent operates in a square gridworld of a given depth $d = \{5, 6, 7\}$. It starts at the top left corner and its goal is to reach the bottom right corner and receive a reward of 1. The agent has two action 1 and 2. While taking action 1 leads the agent downwards and receives no reward, taking the other leads the agent downwards and to the right and receives a reward of -0.001 . Similar to the previous environment, for different depths, we measure the average suboptimality of the agent trained from 10^6 episodes over 30 seeds using 100 different learning rates from $\exp(-9)$ to $\exp(7)$. We observed a similar "bowl" shape across the learning rates. However, the thresholds are different from the previous analysis. Specifically, the learning

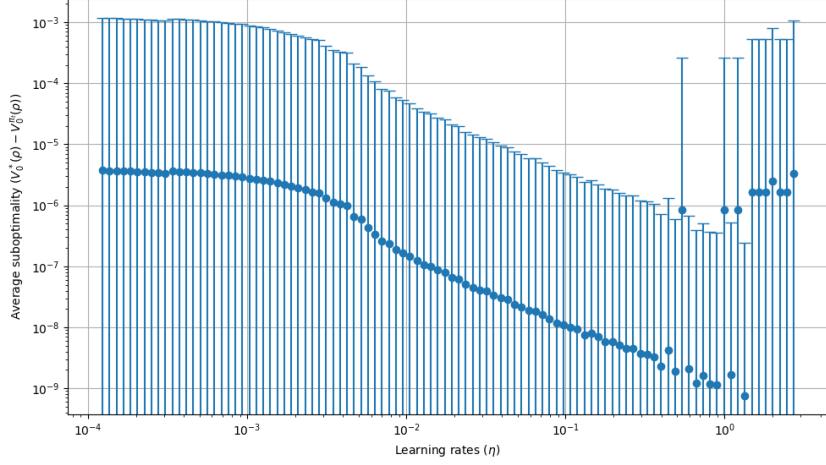


Figure 5: Average last-iterate suboptimality gap of ChainMDP size 4

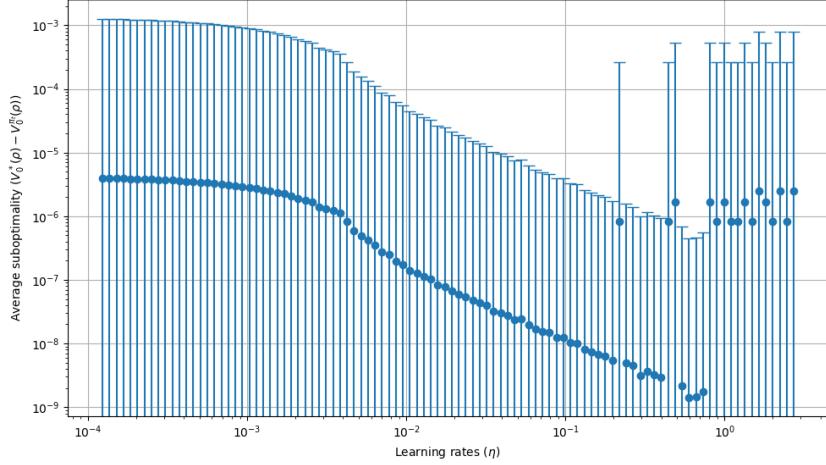


Figure 6: Average last-iterate suboptimality gap of ChainMDP size 5

rate $\eta = 10$ has the lowest suboptimality. Finally, we evaluate the performance of the REINFORCE algorithm (Algorithm 2) in the Cartpole environment. Specifically, we measure the average return received by the agent from 10^5 episodes over 5 seeds using $\eta = \{10^{-5}, 10^{-4}, 10^{-2}, 1\}$. Again, we observed a similar "bowl" shape across learning rates. The learning rate $\eta = 0.01$ achieves the highest average return (approximately 150), while the average return of the others stay around 25. Overall, we consistently find a "bowl-shaped" relationship between the learning rate and performance, and the specific shape and optimal point of this bowl vary significantly with the environment's structure.

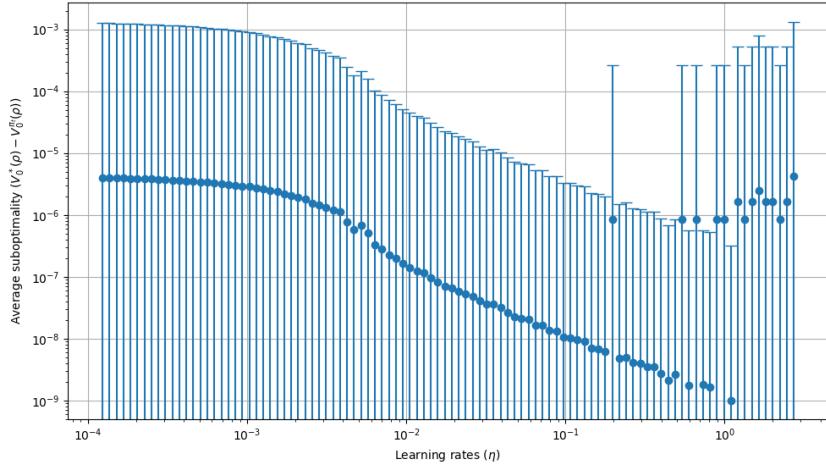


Figure 7: Average last-iterate suboptimality gap of ChainMDP size 6

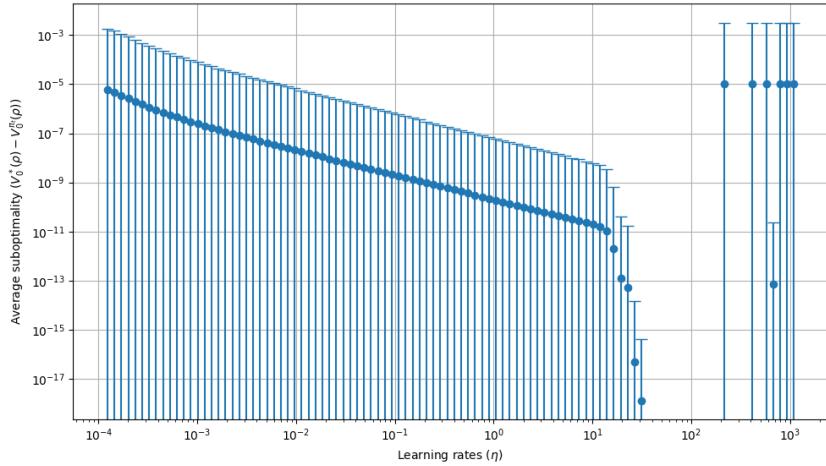


Figure 8: Average last-iterate suboptimality gap of DeepSea depth 5

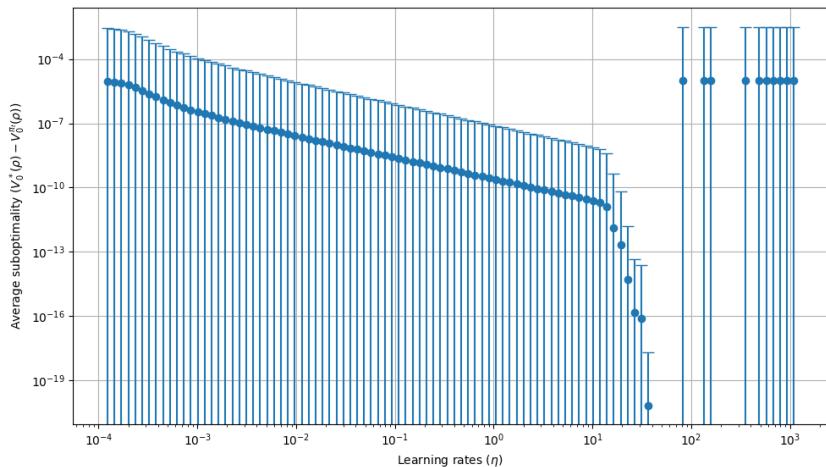


Figure 9: Average last-iterate suboptimality gap of DeepSea depth 6

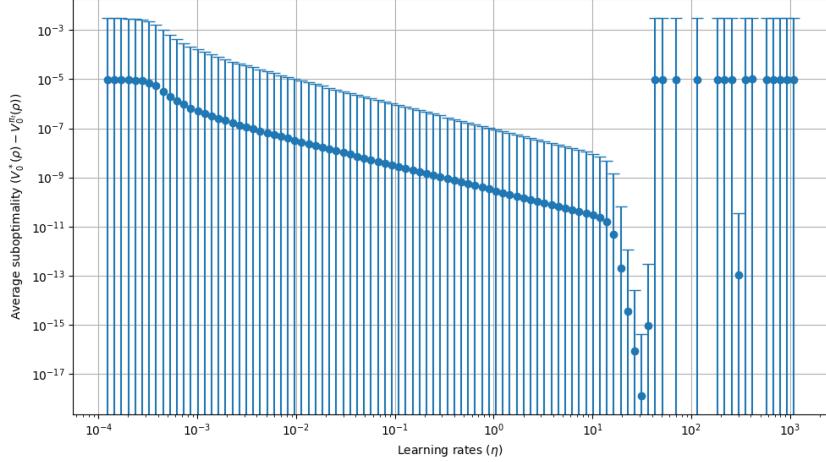


Figure 10: Average last-iterate suboptimality gap of DeepSea depth 7

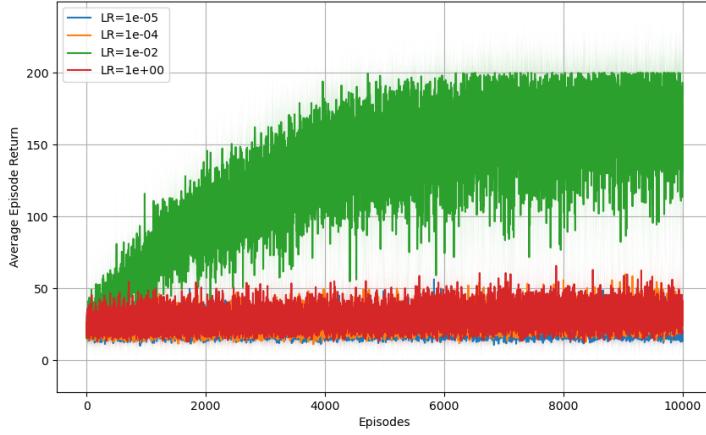


Figure 11: Average suboptimality gap of CartPole

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