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#### Abstract

We present an approach to accelerate statistical inferences in astrophysics by using a combination of neural networks and Hamiltonian Monte Carlo. The neural networks are used to create high-fidelity surrogates of computationally expensive models, while Hamiltonian Monte Carlo accelerates the inferences by more efficiently exploring the parameter space. We demonstrate the potential of this approach by applying it to a realistic model for the Epoch of Reionization.

## 1. Introduction

One of the main objectives of the natural sciences is to formulate models  $m(\theta)$  that explain physical phenomena, and to test their validity by comparing these models with observational data. In astrophysics, this is typically done by obtaining observational data d, and then constraining the model parameters  $\theta$  with Bayes theorem:

$$P(\boldsymbol{\theta}|\boldsymbol{d}) = \frac{\mathcal{L}(\boldsymbol{d}|\boldsymbol{\theta})P(\boldsymbol{\theta})}{P(\boldsymbol{d})}$$
(1)

With this framework, the posterior distribution  $P(\theta|d)$  is calculated given that we assume both a likelihood function  $\mathcal{L}(d|\theta)$ , and a prior distribution  $P(\theta)$ , and that we can calculate the evidence of the model P(d). In most applications, however, P(d) is intractable. Therefore, sampling algorithms designed to approximate the posterior distribution are usually employed. In general, Markov Chain Monte Carlo (MCMC) and Nested Sampling (NS) methods are used for parameter estimation problems in physics and other sciences. The Metropolis-Hastings algorithm (and its variations) is the most commonly used MCMC algorithm (Lewis & Bridle, 2002; Foreman-Mackey et al., 2013; Karamanis et al., 2022), while multi-modal and sliced-based NS algorithms are also widely employed (Feroz & Hobson, 2008; Feroz et al., 2009; Handley et al., 2015).

Unfortunately, applying standard MCMC and NS algorithms can become too computationally expensive to feasibly perform parameter inference. The most obvious problem is that these algorithms typically require multiple model  $m(\theta)$  evaluations. Consequently, complex models that incur in large computational costs can significantly slow down parameter estimation. Furthermore, these algorithms can struggle to efficiently explore high-dimensional parameter spaces, reducing the number of parameters that we are feasibly allowed to vary. To overcome these challenges, we propose using a combination of neural emulators and Hamiltonian Monte Carlo. The structure of this article is follows. In Section 2 we explain the approach of using neural networks alongside Hamiltonian Monte Carlo to accelerate parameter inferences. We then show an application of this method in Section 3. We discuss the results of this application in Section 4 and finish with our conclusions in Section 5.

### 2. The NN + HMC method

#### 2.1. Neural Networks

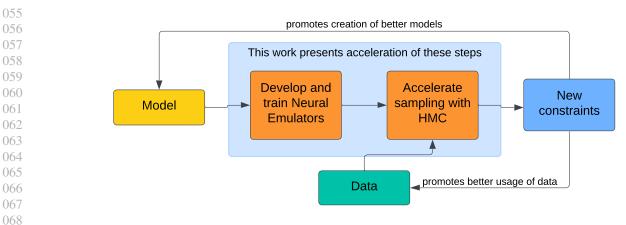
The main goal of an emulator is to provide a high-fidelity approximation of a computationally expensive model that is fast to compute. There are different approaches for creating emulators (Walther et al., 2019; Heitmann et al., 2014; Euclid Collaboration et al., 2019), but neural networks (NN) have become an increasingly popular choice due to their flexibility and ease of use (Piras & Spurio Mancini, 2023; Gong et al., 2023; Nygaard et al., 2023; Ruiz-Zapatero et al., 2024). The specifics of the neural networks (such as the choice of architecture and hyper-parameters) that each application requires may vary depending on the problem itself (see Section 3 for an example). However, we generally expect the neural networks to take the model parameters  $\theta$  as input, and output some model  $m(\theta)$  dependent observable or summary statistic.

#### 2.2. Hamiltonian Monte Carlo

A viable alternative to accelerate sampling is to use gradientbased algorithms. Among these, a popular choice is Hamiltonian Monte Carlo (HMC), a gradient-based modification of the Metropolis-Hastings algorithm. HMC improves the efficiency of the parameter space exploration by proposing

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*Figure 1.* Flow chart that shows the usual workflow of typical inference problems found in astrophysics, emphasizing the challenges that are addressed by employing neural emulators and Hamiltonian Monte Carlo.

samples that have a higher probability of being accepted (Duane et al., 1987). To do so, HMC formulates the problem by creating an analogy between the exploration of the parameter space and the dynamic evolution of a particle living in a phase space defined by the particle's position (chosen to be equal to  $\theta$ ) and momentum p. Typically, this particle is assigned the following Hamiltonian  $\mathcal{H}(\theta, p)$ :

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$$\mathcal{H}(\boldsymbol{\theta}, \boldsymbol{p}) = \frac{1}{2} \boldsymbol{p}^T \mathcal{M}^{-1} \boldsymbol{p} + \boldsymbol{U}(\boldsymbol{\theta})$$
(2)

Where  $\mathcal{M}$  is referred as the mass matrix, and  $U(\theta)$  is the particle's potential energy. The choice of  $U(\theta)$  can depend on the specific application, but the following is generally used:

$$\boldsymbol{U}(\boldsymbol{\theta}) = -ln\big(\mathcal{L}(\boldsymbol{d}|\boldsymbol{\theta})P(\boldsymbol{\theta})\big)$$
(3)

Intuitively, this choice for  $U(\theta)$  causes the regions of high 092 density in the posterior  $P(\theta|d)$  to be located in the regions 093 with low potential energy  $U(\theta)$  (see Equation 1). HMC 094 then increases the efficiency of the standard Metropolis-095 Hastings algorithm by proposing samples that correspond to 096 low energy levels of  $\mathcal{H}(\boldsymbol{\theta}, \boldsymbol{p})$ , which have a higher chance 097 of being accepted. Therefore, the momentum p is randomly 098 sampled from a multivariate Gaussian distribution with zero 099 100 mean and a covariance matrix given by  $\mathcal{M}$  at the beginning of every step of the chain, forcing the hypothetical particle to spend more time exploring the parameter space in regions of low  $U(\theta)$ . It is known that for a target distribution with D dimensions, the cost of an independent sample that HMC 104 produces scales as  $O(D^{5/4})$ , a significant improvement over 105 the scaling of  $O(D^2)$  that the standard Metropolis-Hastings 106 algorithm has (Neal, 2011).

Unfortunately, adopting HMC in most problems is unfea-

sible. This is because the numerical integration that is performed at every step of HMC requires  $\mathcal{L}(d|\theta)$  (and in turn, the model evaluations) to be fully differentiable within the parameter space that is being explored, which is typically not the case in most inference problems. However, replacing the models with neural emulators solves this problem, because we can evaluate their derivatives via auto-differentiation (Rumelhart et al., 1986; Piras & Spurio Mancini, 2023). This allows us to fully take advantage of HMC.

#### 2.3. Pipeline

Figure 1 summarizes the main points of our approach. Since this approach is problem agnostic, creating a flexible pipeline is possible. To this end, we have started the development of a pipeline that facilitates the application of this approach. To ensure that the auto-differentiation evaluations of the networks are efficient, we opt to use JAX as the main backbone for this pipeline. JAX is a high-performance Python package that is mainly designed to work as a machine learning framework (Bradbury et al., 2018). Our pipeline is built to perform the following:

**Training and hyper-parameter tuning:** Given a dataset with pairs of  $(\theta, m(\theta))$ , the pipeline trains the required emulators and performs their hyper-parameter tuning. We use Flax to create the neural networks (Heek et al., 2023), and Optuna to perform the hyper-parameter tuning (Akiba et al., 2019).

**Parameter inference:** Given a set of observations d, the pipeline performs the statistical inference for  $\theta$  using the trained emulators. For the HMC, we use the implementation of the No-U-Turns algorithm (Hoffman & Gelman, 2011) in the Numpyro package (Phan et al., 2019).

**Inference test:** If the ground truth of  $\theta$  is known for *d* (as

is the case with mock observations), then the pipeline is
able to produce a coverage plot that evaluates the validity
of the statistical setup, and allows for a comparison of the
performance of this approach against other methods. See
Section 4.3 for more details on this.

To demonstrate the use of this approach, we apply it in the context of parameter inference for models of the Epoch of Reionization that consider fluctuating ultra-violet backgrounds, as explained in the following section.

#### 3. Application: Constraining the EoR

122 The Epoch of Reionization (EoR) is the period of time in the 123 evolution of the universe during which the neutral hydrogen 124 in the inter-galactic medium (IGM) transitioned from being 125 fully neutral (HI) to fully ionized (HII). A detailed discus-126 sion of the EoR and its study is outside the scope of this paper, but its characterization remains an important chal-128 lenge in modern cosmology (Zaroubi, 2013; Shimabukuro 129 et al., 2023). Among various different probes, the Lyman-130 alpha forest (LAF) is typically used to study the state of 131 the IGM at the later stages of the EoR, as it is a series of 132 redshifted absorption features caused by the presence of HI 133 along the line of sight of observed quasar spectra. 134

135 Unfortunately, modeling the LAF is computationally expen-136 sive, which prohibits a proper quantitative comparison of 137 the wide range of reionization scenarios that still need to be 138 considered. In a typical setup, for instance, high-resolution 139 hydrodynamical simulations are used to model the LAF and 140 its observables for the different stages of the EoR (Doughty 141 et al., 2023). Further post-processing of these simulations is 142 usually done to consider different ultra-violet backgrounds 143 (UVB). The high computational cost of these models makes 144 statistical inferences using standard methods unfeasible, a 145 problem that only worsens with more flexible models that 146 have higher-dimensional parameter spaces. As an initial 147 test of our approach, we replicate some of the results from 148 a previous study (Wolfson et al., 2023), where the LAF is 149 modeled for the case of an inhomogeneous UVB. We will 150 refer to this study as W23 for the remainder of this article. 151 To perform the parameter inference, W23 employs a com-152 bination of Nearest Grid-Point (NGP) interpolation and the 153 standard Metropolis-Hastings algorithm. For this, a grid of 154  $\sim 550$  models of both  $\boldsymbol{\xi}_m$  and  $\boldsymbol{\Sigma}_m$  as a function of  $\lambda_{\mathrm{mfp}}$  and 155  $\langle F \rangle$  for each redshift z that is being considered is generated, 156 and is used to perform inferences on real data. We refer to 157 this approach as the NGP + MH method for the remainder 158 of this article. Replicating the results of W23 with the NN + 159 HMC method serves as a non-trivial test of this technique.

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#### 3.1. Models of EoR

We refer the reader to W23 for a detailed description of the astrophysics and the meaning of the models described in this section. As a brief summary, the model in W23 uses the mean autocorrelation function of the LAF  $\boldsymbol{\xi}_m$  as a summary statistic to constrain two physical parameters related to the EoR: the mean-free path of ionizing photons  $\lambda_{mfn}$ (responsible for ionizing the HI in the IGM), and the mean flux  $\langle F \rangle$  of the flux skewers that are used to calculate  $\boldsymbol{\xi}_m$ . To obtain  $\boldsymbol{\xi}_m$ , a series of hydrodynamical simulations are used to model the LAF. Independent fluctuating UVBs are obtained with a semi-numerical method, and are then used to post-process the output from the simulations. For each simulation, a thousand mock Lyman-alpha flux skewers are forward-modeled, which are then used to calculate individual mock LAF autocorrelation functions  $\xi_i$ . Lastly, the mean autocorrelation function of the LAF  $\boldsymbol{\xi}_m$  is obtained by averaging over all the individual mock autocorrelation functions. Additionally, model-dependent covariance matrices  $\Sigma_m$  are calculated given the statistical setup that is used (see Section 3.2). These are calculated by using the following:

$$\boldsymbol{\Sigma}_{m} = \frac{1}{N_{\text{mocks}}} \sum_{i=1}^{N_{\text{mocks}}} (\boldsymbol{\xi}_{i} - \boldsymbol{\xi}_{m}) (\boldsymbol{\xi}_{i} - \boldsymbol{\xi}_{m})^{T} \qquad (4)$$

Where  $N_{\text{mocks}} = 500000$ . In summary, this model allows us to calculate  $\boldsymbol{\xi}_m$  as a function of  $\lambda_{mfp}$  and  $\langle F \rangle$ , as well as their corresponding, model-dependent covariance matrices  $\boldsymbol{\Sigma}_m$ .

#### 3.2. Parameter inference

To constrain  $\lambda_{mfp}$  and  $\langle F \rangle$  from a mock observation  $\boldsymbol{\xi}_i$ , the following multivariate Gaussian likelihood  $\mathcal{L}(\boldsymbol{\xi}_i | \lambda_{mfp}, \langle F \rangle)$  is assumed:

$$\mathcal{L} = \kappa \exp\left(-\frac{(\boldsymbol{\xi}_i - \boldsymbol{\xi}_m))^T \boldsymbol{\Sigma}_m^{-1}(\boldsymbol{\xi}_i - \boldsymbol{\xi}_m)}{2}\right) \quad (5)$$

Where  $\kappa$  is given by:

$$\kappa = \frac{1}{\sqrt{\det(\boldsymbol{\Sigma}_m)(2\pi)^n}} \tag{6}$$

The proper adoption of this Gaussian likelihood requires a careful consideration of the covariance matrices  $\Sigma_m$  that are used when evaluating the likelihood. Ideally, the covariance matrix is obtained directly from the observed data. With this approach, a noisy approximation of  $\Sigma_m$  is calculated by partitioning the data with bootsrapping or jacknifing, and subsequently recombining the samples. This has the advantage that the obtained estimate of  $\Sigma_m$  is agnostic to

165 any underlying assumptions of the model that is being considered, since it is being determined empirically (Shirasaki 167 et al., 2017; Alam et al., 2017; Hamana et al., 2020). Unfor-168 tunately, there is not enough high-quality data to do this for 169 the autocorrelation function of the LAF at the redshifts in 170 which we are interested. An alternative is to calculate  $\Sigma_m$ 171 from mock observations (Gruen et al., 2015; Krause et al., 172 2017; Martinet et al., 2018), as explained in Section 3.1.

#### 174 3.3. Emulating $\xi_m$ 175

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We first need to construct an emulator for  $\boldsymbol{\xi}_m$ . This emulator 176 needs to take the two input model parameters, and output an emulated mean autocorrelation function of the LAF. We 178 choose to employ a fully connected, feed-forward neural 179 network. For the training, we use the relative root mean 180 squared error (RRMSE) as a loss function. For a single autocorrelation function, the RRMSE is calculated by: 182

$$\text{RRMSE} = \frac{1}{n_b} \sum_{i=1}^{n_b} \frac{\sqrt{(\xi_i - \hat{\xi}_i)^2}}{\xi_i}$$
(7)

Where  $\xi$  is the truth,  $\hat{\xi}$  is its emulation, and  $n_b$  is the number 189 of velocity bins that an individual autocorrelation function 190 has, and over which the mean error is calculated. Even though the training dataset is normalized before training (as 192 is typically done in ML applications), the loss is calculated with the un-scaled autocorrelation functions. This improves the interpretability of the emulation error, as the calculated 195 loss after training gives us a relative error calculated in terms 196 of the original, astrophysical units. 197

#### **3.4.** Emulating $\Sigma_m$ 199

200 Emulating  $\Sigma_m$  requires more careful consideration, as co-201 variance matrices are symmetric. We take advantage of 202 Cholesky decomposition to satisfy this requirement, as it allows us to obtain a lower triangular matrix L (known as a 204 Cholesky factor) for each  $\Sigma_m$ , such that  $\Sigma_m = LL^T$ . To 205 further simplify the emulation, the Cholesky factors are re-206 ordered into 1D arrays that we name the flattened Cholesky factors  $L_{\text{flat}}$ . We opt to also use a fully connected, feed-208 forward neural network to emulate  $L_{\text{flat}}$ , which can then 209 be used to reconstruct an emulated  $\Sigma_m$ . For training, we 210 also use an RRMSE loss (Eq. 7) calculated on the un-scaled 211  $L_{\text{flat}}$  for each covariance matrix in the data set. Figure 5 in 212 Appendix A shows an example of an emulated covariance 213 matrix. 214

#### 4. Results

For the final test, we decided to use a dataset of the models described in Section 3.1 at a redshift of z = 5.1. This

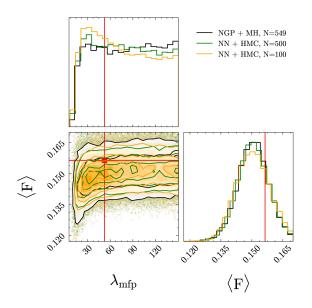


Figure 2. Example of an inference done on a single mock observation. It shows the overlaid corner plot obtained using the three distinct methods: NGP + MH using all 549 models, NN + HMC with  $N_{\text{train}} + N_{\text{val}} = 500$ , and NN + HMC with  $N_{\text{train}} + N_{\text{val}} = 100$ . The red line shows the true parameters of this mock observation, which are  $\lambda_{\rm mfp} = 52.5 \text{Mpc}$  and  $\langle F \rangle = 0.1564$ .

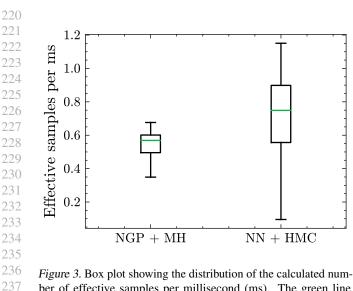
dataset includes  $N_{\text{total}} = 549$  models, which are split into  $N_{\text{train}} = 450 \text{ models for training}, N_{\text{val}} = 50 \text{ for validating},$ and  $N_{\text{test}} = 49$  for testing.

#### **4.1. Emulation performance**

Hyper-parameter tuning was performed for both emulators using Optuna (see Section 2.3) to explore different architecture and training parameters. Table 1 in Appendix B shows the results of the hyper-parameter tuning. We used the average RRMSE calculated on the test set to evaluate the performance of the two emulators. For the  $\boldsymbol{\xi}_m$  emulator, the error was calculated with the un-scaled  $\boldsymbol{\xi}_m$  (the same as during training), and was found to be 0.0034. For the  $\Sigma_m$ , the error was calculated with the covariance matrices (different from training, see Section 3.4), and was found to be 0.016. This means that both emulators produce sub 2% error on average. Figure 6 in Appendix B shows the distribution of these errors for both emulators.

#### 4.2. Inference on mock observations

For comparison purposes, we performed parameter inference on 500 randomly selected mock observations of the autocorrelation function  $\xi_i$  using both the NGP + MH and the NN + HMC methods. For each inference, we set the HMC (MH) to have 16 chains (walkers), taking 3500 steps each and removing 500 steps of each chain (walker) as a



ber of effective samples per millisecond (ms). The green line represents the median, the top and bottom edges of the box represent the first quartile (Q1) and third quartile (Q3) of the data respectively, while whiskers extend from the box to the farthest data point lying within 1.5x the inter-quartile range from the box.

244 burn-in. Figure 2 shows an example of the inference results 245 for a typical model using both of these methods, with two 246 variations of the NN + HMC approach (see Section 4.4). 247 An extensive discussion of the inference results is beyond 248 the scope of this paper, and we refer the reader to W23 for 249 further details. Visually inspecting the results of all 500 250 inferred posterior distributions suggest that both methods 251 produce similar results. 252

As a way to compare the gain in efficiency that using HMC 253 provides, we calculated the number of effective samples 254 per millisecond  $n_{\rm eff}/ms$  using the ArviZ python package 255 (Kumar et al., 2019) for all 500 parameter inferences ob-256 tained with both methods. Figure 3 shows the results of 257 this calculation, showing that the median of  $n_{\rm eff}/ms$  for the 258 NGP + MH and the NN + HMC methods are 0.57 and 0.75259 respectively. Such a marginal gain in efficiency is to be expected given that the problem that we have tested it on 261 has a small dimensional space (2 parameters), but we expect this gain to be more significant as the number of parameters 263 increases (see Section 2.2). 264

#### 4.3. Inference test

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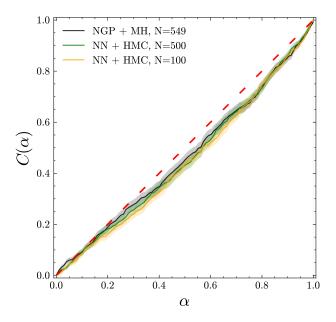
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> To further validate our results, we perform an inference test by creating coverage plots for both methods. Coverage plots are used to test the fidelity of the inference framework that is being employed, and it allows to test the validity of any assumptions that are done during inference (Prangle et al., 2013; Sellentin & Starck, 2019). M23 found that this statistical setup (see Section 3.2) leads to overconfident

posteriors, which is likely caused by the assumption of a Gaussian likelihood. Since we performed all the inferences on mock observations (where we know the ground truth), it is possible for us to make a direct comparison of the coverage that our method produces. Figure 4 shows a comparison of the coverage obtained with the NGP + MH and the NN + HMC methods. As we can see, our method is able to reproduce the same coverage, providing strong evidence that the neural emulators and their application with HMC are producing posterior distributions that have similar statistical behavior to the previous method.



*Figure 4.* Comparison of coverage plots between different methods. We can see that using the full dataset for training the emulators produces similar results to the original method from M23, and that using a reduced dataset produces competitvely similar results (with a small degradation), at a significant decrease of computational cost.

#### 4.4. Reducing the training set

As a last test to our method, we experimented with decreasing the size of the training and validation data sets to see how the inference would perform. For this, a reduced training and validation sets with  $N_{\text{test}} = 65$  and  $N_{\text{val}} = 35$  were created by randomly selecting from the original sets. After hyper-parameter tuning, the performance of the emulators was tested with the same testing set as in Section 4.1. Both emulators obtained a sub 2% error on average, with the  $\xi_m$ emulator obtaining an RMMSE of .0046, and the  $\Sigma_m$  an error of 0.014. The inference on the same 500 mocks was then performed using this new emulators for comparison, and a coverage plot was also produced. Figure 2 shows the 275 results of doing the inference with these emulators on a typ-276 ical mock, and Figure 4 shows the coverage plot. As we can 277 observe, there is a noticeable degradation in performance 278 in the inference, likely caused by having significantly less 279 models in the training and validation sets. However, the re-280 sults are competitive, and provide evidence on how the NN 281 + HMC method can further reduce the overall computational 282 cost of this inference task by requiring significantly fewer 283 models.

# 285286**5. Conclusions**

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287 We started the development of a pipeline that is able to apply the NN + HMC method and then tested it in the context 289 of cosmological parameter inference. We demonstrated 290 the potential that this method has to accelerate statistical 291 inferences while also reducing the overall computational 292 cost. In future work, we will continue the development of 293 our software, and apply it to a more complex problem with 294 a higher dimensional parameter space that will better show 295 the advantages of this approach. 296

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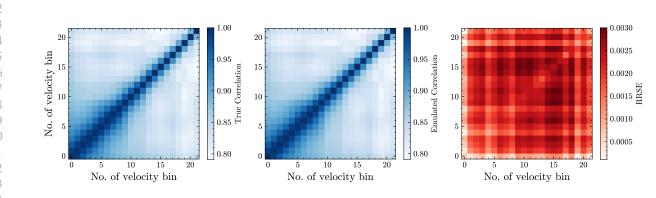
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## A. Example of Covariance Matrix Emulation



*Figure 5.* Example of an emulation of a covariance matrix, visualized as correlation matrices. The left correlation matrix represents the true covariance matrix, the middle one shows the corresponding emulated covariance matrix, and the right shows a color map of the relative error.

As explained in Section 3.4, emulating covariance matrices require special attention given that they are symmetric. Taking advantage of Cholesky decomposition and reordering the resulting lower triangular matrices into 1D arrays allows us to employ a simple architecture for the neural network. To visualize a covariance matrix  $\Sigma$ , we first convert it into a correlation matrix C using the following:

$${C}_{i,j} = rac{{oldsymbol{\Sigma}}_{i,j}}{\sqrt{{oldsymbol{\Sigma}}_{i,i}{oldsymbol{\Sigma}}_{j,j}}}$$

Where the i and j subscripts refer the element in the ith and jth element in the covariance matrix. Figure 5 shows an example of a typical covariance matrix from the model explained in Section 3.1, its corresponding emulated covariance matrix, and the emulation error.

#### B. Performance of emulators

#### B.1. Hyper-parameter tuning

The hyper-parameter tuning is carried out with Optuna, as explained in Section 2.3. For the application in Section 3, the hyper-parameters that were varied for both emulators were the number of hidden layers, the number of perceptrons in each hidden layer, the number of epochs, and the starting learning rate (for this application, both emulators were trained with a cosine scheduler for the learning rate). Table 1 shows the results of the hyper-parameter tuning for both emulators.

Table 1. Results of the hyper parameter-tuning for both emulators.

EMULATOR	LAYERS	Perceptrons	STARTING LEARNING RATE	EPOCHS
$egin{array}{c} oldsymbol{\xi}_m \ oldsymbol{\Sigma}_m \end{array}$	5	[4, 8, 12, 16, 20]	0.03523	750
	6	[4, 8, 8, 8, 16, 20]	0.0393	1500

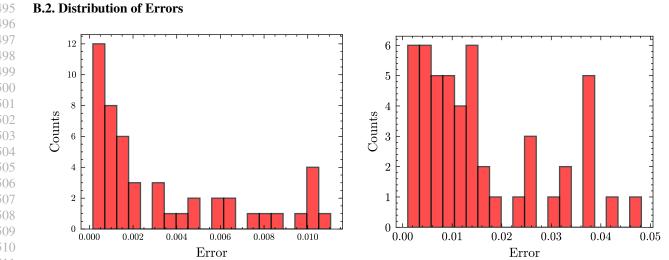


Figure 6. Distribution of errors calculated on the test set ( $N_{\text{test}} = 49$ ) to evaluate the performance of the emulators trained with  $N_{train} + N_{val} = 500$ . Left: Distribution of RRMSE calculated with the  $\boldsymbol{\xi}_m$  emulator. Right: Distribution of RRMSE calculated for the  $\boldsymbol{\Sigma}_m$  emulator.