Cheap and Deterministic Inference for Deep State-Space Models of Interacting Dynamical Systems

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Abstract

Graph neural networks are often used to model interacting dynamical systems since they gracefully scale to systems with a varying and high number of agents. While there has been much progress made for deterministic interacting systems, modeling is much more challenging for stochastic systems in which one is interested in obtaining a predictive distribution over future trajectories. Existing methods are either computationally slow since they rely on Monte Carlo sampling or make simplifying assumptions such that the predictive distribution is unimodal. In this work, we present a deep state-space model which employs graph neural networks in order to model the underlying interacting dynamical system. The predictive distribution is multimodal and has the form of a Gaussian mixture model, where the moments of the Gaussian components can be computed via deterministic moment matching rules. Our moment matching scheme can be exploited for sample-free inference leading to more efficient and stable training compared to Monte Carlo alternatives. Furthermore, we propose structured approximations to the covariance matrices of the Gaussian components in order to scale up to systems with many agents. We benchmark our novel framework on two challenging autonomous driving datasets. Both confirm the benefits of our method compared to state-of-the-art methods. We further demonstrate the usefulness of our individual contributions in a carefully designed ablation study and provide a detailed empirical runtime analysis of our proposed covariance approximations.

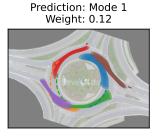
1 Introduction

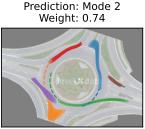
Many dynamical systems such as traffic flow (Li et al., 2018; Yu et al., 2018), fluid dynamics (Ummenhofer et al., 2019) or human motion (Jain et al., 2016) involve interactions between agents. *Graph Neural Networks* (GNN) (Battaglia et al., 2018) have recently emerged as a powerful tool in these settings since they allow to learn the dynamics of interacting systems from data only. For deterministic systems, such as complex physical simulators, recent research (Sanchez-Gonzalez et al., 2020) has made great advances by being able to extrapolate from systems with a small number of agents and short time horizons to systems with a high number of agents and long time horizons.

However, for many real-world applications, predicting a single future trajectory for each agent is not enough, since the stochasticity in the dynamical system has significant consequences. For instance, in autonomous driving, the driver's intention (e.g. overtaking, turning, lane changing) is a hidden factor that may induce different modes of driving trajectories.

There are two established model families that can account for model uncertainty in dynamical systems: a) recurrent methods that accumulate uncertainty over time by applying a fixed transition function repeatedly at each step (Yang et al., 2020; Ivanovic & Pavone, 2019), and b) history-based methods that directly output the predictive distribution for a fixed time horizon over future trajectories (Casas et al., 2020; Mohamed et al., 2020). Predicting the output distribution in a recurrent manner respects the causal order of the dynamical system, i.e. the distribution of time point t is needed in order to compute the distribution at the next time point t+1. However, propagating a distribution through a non-linear recurrent system, as it is required in this case, cannot be performed in closed-form, and existing methods build on *Monte Carlo* (MC) simulations.

Ground Truth





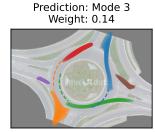


Figure 1: We approximate the predictive distribution of a latent *Graph Deep State-Space Model* (GDSSM) as a Gaussian mixture distribution via deterministic moment matching rules. Given historical information in the form of observed trajectories (dashed lines), our proposed GDSSM architecture predicts the future dynamics while taking interactions between traffic participants into account. We present as solid lines the true future trajectories (left-most plot) and show for each mode and traffic participant, the predicted 95% confidence interval of our model (three right-most plots). Our model accounts for interactions, for example, the brown vehicle is only entering the roundabout if the blue vehicle is staying in the roundabout (Mode 1). If the blue vehicle is leaving the roundabout, the entering lane for the brown vehicle is blocked by the blue vehicle (Mode 2, Mode 3) and the brown car has to wait. The ground truth data as well as the map is part of the rounD dataset (Krajewski et al., 2020).

In interacting systems, the sample space grows linearly with the number of agents, requiring a high number of MC samples which can make these methods prohibitively slow. In contrast, history-based methods directly predict the distribution over future trajectories mitigating the sampling overhead. However, this approach makes the learning problem hard as the model needs to learn the future distribution for multiple time steps ahead. In order to account for this increasing model complexity, large network architectures are necessary that can restrict its usage on embedded systems with limited memory capacity.

In this work, we present a novel approach for modeling stochastic dynamical systems with interacting agents that is able to generate expressive multi-modal predictive distributions over future trajectories in a runtime and memory efficient manner. We model the unknown stochastic dynamical system as a Deep State-Space Model (DSSM) in which the shared dynamics of all agents are modeled in a joint latent space using GNNs. Our model belongs to the family of recurrent neural networks and we replace the expensive MC operations during training and inference by introducing a novel deterministic moment matching scheme. Prior work (Look et al., 2020) on moment matching for dynamical systems does not consider interacting systems and is restricted to unimodal processes. We overcome the first limitation by applying GNNs in the transition model and the second limitation by placing a Gaussian Mixture Model (GMM) over the initial latent state. For each mixture component, we independently apply our moment matching rules in order to arrive at multimodal predictive distributions over future trajectories. In autonomous driving, the initial latent state can be estimated from historical information and is closely linked to the drivers' intentions. Given the intentions, the predictive distribution can often be accurately modeled with an unimodal distribution (Cui et al., 2019; Chai et al., 2019). Finally, as there exists a wide variety of dense traffic scenarios, the high number of agents can result in prohibitively large GMM covariance matrices as their size grows quadratically with the number of traffic participants. We address this problem by proposing structured covariance approximations.

We summarize our contribution as follows:

- We derive output moments for GNN layers, which makes GNNs applicable to moment matching algorithms leading to the first deterministic inference scheme for deep state-space models for interacting systems.
- We introduce a GMM distribution over the initial latent states that results in multimodal predictive distributions over future trajectories.

• We propose structured approximations to the GMM covariance matrices that can reduce the computational complexity of our approach from $O(M^3)$ to $O(M^2)$, where M is the number of agents.

In our experiments, we benchmark our proposed model on two challenging autonomous driving datasets. Our results demonstrate that our deterministic model produces accurate and well-calibrated predictions compared to state-of-the-art alternatives. We visualize the predictive output distribution for a real-world traffic scenario on a roundabout with multiple agents in Fig. 1. The future distribution is highly multimodal, as traffic participants can leave the roundabout from several exits. Our model is capable of predicting multiple modes, which we efficiently approximate as a GMM and takes interaction into account by using GNNs.

To gain further insights into our model and inference scheme, we carefully examine the impact of the individual contributions of our work in an ablation study. Finally, we provide an empirical runtime study of our covariance approximations. Our findings indicate that sparse covariance approximations are favourable for applications with low computational resources as they reduce the computational complexity up to a factor of 100 and we analyse its implications on accuracy and calibration in detail.

2 Background

In this chapter, we provide background on (deep) state-space models for single-agent systems and on graph neural networks for interaction modeling. Both together form the basis for our new model for stochastic dynamical interacting systems which we introduce in Sec. 3.

2.1 Deep State-Space Models

State Space Models (SSM) are a model class for dynamical systems (e.g. Schön et al. (2011), Särkkä (2013)) that assume that each D_y -dimensional observed variable $y_t \in \mathbb{R}^{D_y}$ is emitted by a latent D_x -dimensional latent variable $x_t \in \mathbb{R}^{D_x}$. The latents are coupled via first-order Markovian dynamics, e.g. the state at time point x_t only depends on the state of the previous time point x_{t-1} . Typically the observed state y_t does not contain all necessary information in order to reliably predict the next observed state y_t . Consider for example the case of traffic forecasting in which the observed state y_t only contains the position of a vehicle but is missing velocity and acceleration data. We can then accommodate the latent state x_t with the missing information in order to allow for accurate forecasts about the next time point. Consequently, SSMs are a flexible model class that allows us to make reliable forecasts about complex systems.

A Deep State-Space Model (DSSM) is a non-linear SSM in which the transition model, that maps the latent state from the current to the next time point, and the emission model, that maps the latent state to the outputs, are realized by neural networks. They come in handy for applications in which the true underlying dynamics are not known and must be estimated from data. Assuming additive Gaussian noise (e.g. Krishnan et al. (2017)), their generative model can be written down as follows

$$x_0 \sim p(x_0|\mathcal{I}),\tag{1}$$

$$x_t \sim \mathcal{N}(x_t | x_{t-1} + f(x_{t-1}, \mathcal{I}), \operatorname{diag}(L(x_{t-1}, \mathcal{I}))), \qquad t = 1, \dots, T$$
 (2)

$$y_t \sim \mathcal{N}\left(y_t | g(x_t), \operatorname{diag}\left(\Gamma(x_t)\right)\right), \qquad t = 0, \dots, T$$
 (3)

where $\mathcal{I} \in \mathbb{R}^{D_{\mathcal{I}}}$ is the context variable that encodes auxiliary information, such as historical or relational information. The mean update $f(x_t, \mathcal{I}) : \mathbb{R}^{D_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{D_x}$, governing the deterministic component of the transition model, is parameterized by a neural net with an arbitrary architecture. Similarly, the variance update $L(x_t, \mathcal{I}) : \mathbb{R}^{D_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{D_x}$ is parameterized by another neural net, which models the stochasticity of the system. Both f and L are neural networks, which are parameterized by $\theta = \{\theta_f, \theta_L\}$. The emission model follows a Gaussian distribution with mean $g(x_t) : \mathbb{R}^{D_x} \to \mathbb{R}^{D_y}$ and variance $\Gamma(x_t) : \mathbb{R}^{D_x} \to \mathbb{R}^{D_y}$, where g and Γ are both neural networks with arbitrary architecture and parameters $\psi = \{\psi_g, \psi_{\Gamma}\}$.

Assuming additive Gaussian noise allows us to interpret the transition model [Eq. (2)] as a discretized neural stochastic differential equation (Tzen & Raginsky, 2019; Look et al., 2020). While we do not pursue this line

of work any further, we note that this connection allows for straight-forward extensions to irregular sampled time series. Finally, there exists also work that couples state-space models with recurrent neural networks (Chung et al., 2015; Fraccaro et al., 2016) whose gating mechanism can help in learning long-term effects.

2.2 Graph Neural Networks

Graph neural networks (GNNs) have emerged as a powerful method for interaction modeling (Battaglia et al., 2018; Hamilton et al., 2017; Gilmer et al., 2017). Given a set of agents and relational information in form of a graph, each agent corresponds to one node in the graph that is equipped with a set of features. The relation between the agents is encoded via the edges and information exchange between the agents takes place by sending messages along the edges. By performing multiple rounds of message-passing, information can flow along the graph and allows for interactions between non-adjacent agents provided that a path between the agents exists.

More formally, we define the structure of our GNN as follows. For M agents, a GNN receives as inputs a set of node features $x = \{x^m\}_{m=1}^M$, where $x \in \mathbb{R}^{MD_x}$ and $x^m \in \mathbb{R}^{D_x}$, and a set of edges $\mathcal{E} = \{e^{m,m'}\}_{m,m'=1}^M$ which is part of the context variable $\mathcal{I} \in \mathbb{R}^{D_{\mathcal{I}}}$. The edge attribute $e^{m,m'}$ has a binary encoding, where $e^{m,m'} = 1$ if agent m and agent m' are related. The GNN output is an update of the node features, i.e. $z = \text{GNN}(x,\mathcal{I})$ with $z \in \mathbb{R}^{MD_z}$, and consists of the following two steps:

1. For each agent m, receive message $x^{\mathcal{N}_m} \in \mathbb{R}^{D_x}$ by aggregating information from neighboring agents:

$$x^{\mathcal{N}_m} = \operatorname{AGG}\left(\left\{x^{m'}|e^{m,m'} = 1\right\}\right) \in \mathbb{R}^{D_x},\tag{4}$$

where $\{x^{m'}|e^{m,m'}=1\}$ denotes the set of all neighbours of node m. The aggregation operation AGG is permutation invariant, i.e. it does not change when the ordering of the inputs is swapped and generalizes to a varying number of inputs. A commonly used aggregation operation that we also apply in our work is the mean function.

2. For each agent m, update the node information:

$$z^{m} = \text{UPDATE}(x^{m}, x^{\mathcal{N}_{m}}, \mathcal{I}), \tag{5}$$

where UPDATE $(x^m, x^{\mathcal{N}_m}, \mathcal{I}): \mathbb{R}^{D_x} \times \mathbb{R}^{D_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{D_z}$ is typically implemented by a neural network.

A simple form of an interacting dynamical system takes the features of each agent at its current position and connects agents that are within a pre-defined radius with edges. The GNN operation updates the position and velocity information of each agent by taking information of the adjacent traffic participants into account.

3 Graph Deep State-Space Models

We aim to model stochastic dynamical interactions between agents following complex behavioural patterns, such as road traffic interactions. We extend deep state-space models to interacting systems by proposing *Graph Deep State-Space Models* (GDSSM) by employing graph neural networks in the transition model in order to efficiently model interactions between agents. After having defined our probabilistic model in this section, we will introduce in the subsequent section a novel scheme for efficient and deterministic training and predictions.

We are interested in modeling the dynamics of M interacting agents with deep state-space models by using a coupled latent space. In other words, instead of using a D_x -dimensional latent space for each agent, we assume that the agents share a latent space of size MD_x . Since (i) the number of agents can vary between scenes, (ii) the transition model should be agnostic to the order of the agents and (iii) it is challenging to parameterize high-dimensional latent spaces, we opt for using GNNs in the transition model.

More formally, we denote the state of agent m at time step t as $x_t^m \in \mathbb{R}^{D_x}$ and the set of all state variables as $x_t = \{x_t^m\}_{m=1}^M$. The dynamics of x_t follow Eq. (2), where the mean update $f(x_t, \mathcal{I}) : \mathbb{R}^{MD_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{MD_x}$ and the variance update $L(x_t, \mathcal{I}) : \mathbb{R}^{MD_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{MD_x}$ are implemented with the help of graph neural networks

$$f(x_t, \mathcal{I}) = \begin{bmatrix} \tilde{f}(x_t^1, x_t^{\mathcal{N}_1}, \mathcal{I}) \\ \vdots \\ \tilde{f}(x_t^M, x_t^{\mathcal{N}_M}, \mathcal{I}) \end{bmatrix}, \qquad L(x_t, \mathcal{I}) = \begin{bmatrix} \tilde{L}(x_t^1, x_t^{\mathcal{N}_1}, \mathcal{I}) \\ \vdots \\ \tilde{L}(x_t^M, x_t^{\mathcal{N}_M}, \mathcal{I}) \end{bmatrix}.$$
(6)

The agent-specific mean update is denoted by $\tilde{f}(x_t^m, x_t^{\mathcal{N}_m}, \mathcal{I}) : \mathbb{R}^{D_x} \times \mathbb{R}^{D_x} \times \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{D_x}$ and the variance by $\tilde{L}(x_t^m, x_t^{\mathcal{N}_m}, \mathcal{I}) : \mathbb{R}^{D_x} \times \mathbb{R}^{D_x} \times \mathbb{R}^{D_x} \times \mathbb{R}^{D_x} \times \mathbb{R}^{D_x}$. Both implement the update function in general graph neural networks [Eq. (5)], whereas $, x_t^{\mathcal{N}_m}$ contains aggregated information of the states from all neighboring agents [Eq.(4)]. A deterministic variant of our model, e.g. setting $L(x_t, \mathcal{I}) = 0$, has been successfully used for learning surrogate models for complex physical systems (Sanchez-Gonzalez et al., 2020). We further assume that it is sufficient to couple the latent dynamics across the agents and keep the emission model [Eq. (3)] independent across agents.

Furthermore, we note that although the transition noise factorizes across agents, our model is capable of modeling correlations between agents since the mean and the variance depend not only on the state of the m-th agent, but also on the states of all neighboring agents. In consequence, after a aggregation steps, our model accounts for correlations between agent m and agent m' provided that they are connected by a path that is at most a steps long.

In order to complete the probabilistic description of our model, we further specify the distribution of the initial latent state $x_0 \in \mathbb{R}^{MD_x}$ with a Gaussian Mixture Model (GMM)

$$p(x_0|\mathcal{I}) \sim \sum_{v=1}^{V} \pi_v(\mathcal{I}) \mathcal{N}(\mu_{0,v}(\mathcal{I}), \operatorname{diag}(\Sigma_{0,v}(\mathcal{I}))),$$
 (7)

where V is the number of mixture components. Each component is specified by its weight $\pi_v(\mathcal{I}): \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}_+$, mean $\mu_{0,v}(\mathcal{I}): \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{MD_x}$, and diagonal covariance $\Sigma_{0,v}(\mathcal{I}): \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{MD_x}_+$. The weights $\pi_{1:V}$ form a standard V-simplex. We use a GNN, which we refer to as the embedding function $h(\mathcal{I}): \mathbb{R}^{D_{\mathcal{I}}} \to \mathbb{R}^{V+2VMD_x}$, in order to model the initial state distribution

$$h(\mathcal{I}) = \begin{bmatrix} \pi_{1:V}(\mathcal{I}) \\ \mu_{0,1:V}(\mathcal{I}) \\ \Sigma_{0,1:V}(\mathcal{I}) \end{bmatrix}. \tag{8}$$

We assume that the context variable \mathcal{I} contains relational information as a set of edges as well as historical information for each agent in the form of an observed trajectory. In a sense, the embedding function acts hereby as a filter, which learns a distribution over the initial latent state from past observations.

In autonomous driving, the initial latent state encodes the drivers' intention. The context information \mathcal{I} is often not sufficient to rule out different hypotheses about the future, e.g. does the car behind us want to overtake in the next five seconds or not. Using a mixture model allows us to incorporate different hypotheses into the model in a principled manner which will ultimately lead to highly multimodal predictive distributions.

State-space models and graph neural networks have been previously combined for multi-agent trajectory forecasting in Yang et al. (2020). In contrast to our work, the authors (i) use a slightly different model definition by applying recurrent neural networks and a non-Gaussian density in the transition model and (ii) perform Monte Carlo sampling during inference which can lead to slow convergence. In the next chapter, we show that our model definition allows for more efficient training by performing deterministic moment matching rules. We compare to Monte Carlo alternatives in our experiments.

4 Deterministic Approximations for GDSSMs

In this chapter, we present our novel inference scheme for GDSSMs that enables efficient and deterministic training and predictions. In Sec. 4.1, we first give a short overview over existing inference techniques and compare it to our scheme which aims at directly maximizing the predictive log-likelihood of future trajectories. Core to our algorithm is the so-called transition kernel, i.e. $p(x_t|x_0,\mathcal{I})$ which allows to propagate the latent state forward in time. The transition kernel can be computed deterministically by using *Bidimensional Moment Matching* (BMM) (Look et al., 2020). We review its core ideas in Sec. 4.2 and propose an efficient approximation to the predictive log-likelihood based on it in Sec. 4.3. Our moment matching rules necessitate the computation of output moments and expected Jacobians of graph neural network layers. We present output moments for commonly used layers in Sec. 4.4. As our algorithm approximates the output distribution at each time step with a Gaussian mixture distribution over all agents, the resulting covariance matrix for each mixture component can become computationally intractable for a large number of traffic participants. We address this pain point in Sec. 4.5 by proposing sparse approximations to the covariance matrix.

4.1 Inference

Classical inference methods for state-space models aim at directly maximizing the log-likelihood of the data

$$\log p(y_0, \dots, y_T | \mathcal{I}) = \log \int p(x_0 | \mathcal{I}) p(y_0 | x_0) \prod_{t=1}^T p(x_t | x_{t-1}, \mathcal{I}) p(y_t | x_t) dx_0 \dots dx_T, \tag{9}$$

where $p(x_t|x_{t-1})$ is defined in Eq. (2) and $p(y_t|x_t)$ in Eq. (3). This quantity can only be computed in closed-form if the emission and transition model are linear Gaussians. In our case, the transition model is parameterized by a graph neural network and the emission model by a standard neural network. Both functions are highly non-linear and render an analytical solution to Eq. (9) infeasible. Therefore, existing methods apply either a particle filter (Schön et al., 2015), variational inference (Krishnan et al., 2017; Bayer et al., 2021) or a combination of both (Naesseth et al., 2018) in order to approximate the log-likelihood. All of these approaches have in common that a good approximation to the smoothing distribution, i.e. $p(x_0, \ldots, x_T|y_0, \ldots, y_T, \mathcal{I})$, is central for the algorithm to succeed.

However, the smoothing distribution is only used as an auxiliary tool during inference. For many prediction tasks, the quantity of interest is the *predictive log-likelihood*

$$PLL(y_1, \dots, y_T | \mathcal{I}) = \sum_{t=1}^{T} \log p(y_t | \mathcal{I})$$
(10)

$$= \sum_{t=1}^{T} \log \int p(x_0|\mathcal{I}) p(x_t|x_0, \mathcal{I}) p(y_t|x_t) dx_0 dx_t, \tag{11}$$

where the transition kernel $p(x_t|x_0,\mathcal{I})$ in the above equation is available via the recurrence

$$p(x_t|x_0, \mathcal{I}) = \int p(x_t|x_{t-1}, \mathcal{I})p(x_{t-1}|x_0, \mathcal{I})dx_{t-1}$$
(12)

and $p(x_0|\mathcal{I})$ is given by Eq. (7). In contrast to the standard training objective, the predictive log-likelihood propagates the latent state forward in time without receiving any feedback from the observations mimicking the behavior during prediction time. Since we want to use the same objective during training and test time, we opt for directly maximizing the predictive log-likelihood during training. Note that a similar argumentation has also been made in Bengio et al. (2015) to generate a sequences of tokens.

It is worth noting that the t-step transition kernel $p(x_t|x_0,\mathcal{I})$ with t>1 cannot be computed in closedform since the distribution $p(x_{t-1}|x_0,\mathcal{I})$ has to be propagated through the non-linear transition model $p(x_t|x_{t-1},\mathcal{I})$. In the past, various approximations have been proposed that can be roughly split into two groups: (a) MC sampling based approaches (Brandt & Santa-Clara, 2002; Pedersen, 1995; Elerian et al., 2001) and (b) deterministic approximations based on assumed densities (Särkkä et al., 2015). While MC based approaches can, in the limit of infinite many samples, approximate arbitrarily complex distributions, they are often slow in practice and their convergence is difficult to assess. In contrast, deterministic approaches often make the simplifying assumption that the t-step transition kernel can be approximated by a Gaussian distribution. This assumption can be justified if a good locally linear approximation to the transition model exists and the observations are densely sampled. An assumed density approach tailored towards neural Stochastic Differential Equations (SDEs) was proposed in Look et al. (2020), in which the authors discretized the differential equations and then approximated the transition kernel using a Gaussian density by performing Bidimensional Moment Matching (BMM) in time direction and across neural network layers. Since the transition model in SSMs can be interpreted as discretized SDEs (Särkkä et al., 2015), we build in our work on their approach as it is computationally efficient and leads to stable training.

4.2 Bidimensional Moment Matching

In order to compute the transition kernel we need to solve the nested set of integrals described in Eq. (12), which is in general not possible in closed form. A common approximation is to assume a Gaussian transition kernel for all time steps and then propagate its moments along time direction using a numerical integration scheme (Särkkä & Sarmavuori, 2013; Särkkä et al., 2015). Prior work in the context of neural SDEs (Look et al., 2020) proposes a deterministic approximation of the transition kernel, which relies on numerical integration via moment propagation through neural network layers. This approach was shown to be superior over standard numerical integration schemes in terms of compute and accuracy, and forms the basis of our algorithm. Moment propagation through neural network layers is also used along similar lines in the context of expectation propagation (Hernandez-Lobato & Adams, 2015; Ghosh et al., 2016), deterministic variational inference (Wu et al., 2019), and evidential deep learning (Haussmann et al., 2020).

In this section, we recapitulate the original BMM algorithm of Look et al. (2020) and its use for propagating the latent state forward in time [Eq. (2)]. In Sec. 4.3, we show how the algorithm can be extended to state-space models and multimodal predictions, while in Sec. 4.4 we derive moment matching rules for graph neural network layers.

BMM approximates the transition kernel $p(x_t|x_0,\mathcal{I})$ by combining horizontal moment matching along the time axis with vertical moment matching across the neural network layers.

Horizontal Moment Matching In order to facilitate the computation of the transition kernel we replace $p(x_t|x_0,\mathcal{I})$ for all time steps $t=1,\ldots,T$ with a Gaussian distribution

$$p(x_t|x_0, \mathcal{I}) = \int p(x_t|x_{t-1}, \mathcal{I})p(x_{t-1}|x_0, \mathcal{I})dx_{t-1},$$

$$\approx \mathcal{N}(x_t|\mu_t(\mathcal{I}), \Sigma_t(\mathcal{I})),$$
(13)

with mean $\mu_t(\mathcal{I})$ and covariance $\Sigma_t(\mathcal{I})$. This approximation simplifies the problem to calculating the first two moments of the transition kernel and is assumed to work well if the dynamics can be locally approximated by a linear model, which is the case for many applications. Mean μ_t and covariance Σ_t are available as a function of prior moments μ_{t-1} and Σ_{t-1} (Look et al., 2020)

$$\mu_t(\mathcal{I}) = \mu_{t-1}(\mathcal{I}) + \mathbb{E}[f(x_{t-1}, \mathcal{I})] \tag{14}$$

$$\Sigma_{t}(\mathcal{I}) = \Sigma_{t-1}(\mathcal{I}) + \text{Cov}[f(x_{t-1}, \mathcal{I})] + \text{Cov}[x_{t-1}, f(x_{t-1}, \mathcal{I})] + \text{Cov}[x_{t-1}, f(x_{t-1}, \mathcal{I})]^{T} + \text{diag}\left(\mathbb{E}[L(x_{t-1}, \mathcal{I})]\right),$$

where $\text{Cov}[x_{t-1}, f(x_{t-1}, \mathcal{I})]$ denotes the cross-covariance between the random vectors in the arguments.

Vertical Moment Matching Mean $\mathbb{E}[f(x_{t-1},\mathcal{I})]$ and covariance $\text{Cov}[f_{\theta}(x_{t-1},\mathcal{I})]$ of the transition function, as well as the expected variance update $\mathbb{E}[L(x_{t-1},\mathcal{I})]$ can be computed as a result of moment propagation through neural network layers. For many common layers, including affine transformations and ReLU activation functions, the corresponding output moments can be either computed in closed-form or

good approximations are available in the literature (Wu et al., 2019). In contrast, approximating the cross-covariance $\text{Cov}[x_t, f_{\theta}(x_t, \mathcal{I})]$ cannot be achieved using moment matching rules since we cannot decompose the cross-covariance term into layerwise operations. Instead, we resort to Stein's Lemma using

$$Cov[x_t, f(x_t, \mathcal{I})] = Cov[x_t] \mathbb{E}[\nabla_{x_t} f(x_t, \mathcal{I})], \tag{15}$$

where the expected Jacobian can be approximated as (Look et al., 2020)

$$\mathbb{E}[\nabla_{x_t} f(x_t, \mathcal{I})] \approx \prod_{l=1}^{L} \mathbb{E}[J_t^l]. \tag{16}$$

Above, J_t^l denotes the Jacobian at layer l at time step t. The expectation of the Jacobian is analytically available or can be closely approximated for common layer types.

4.3 Approximating the Predictive Log-Likelihood

We are interested in the predictive log-likelihood $\operatorname{PLL}(y_1,\ldots,y_T|\mathcal{I})$ which describes the predictive log-likelihood of all traffic participants up to time step T. In order to calculate it, we need to solve the nested set of integrals given in Eq. (11). The BMM algorithm allows us to approximate the transition kernel $p(x_t|\mathcal{I}) = \int p(x_t|x_0,\mathcal{I})p(x_0|\mathcal{I})dx_0$ in case that the initial state x_0 is a Gaussian distribution. However, in our model formulation the initial latent state x_0 follows a GMM to allow for multimodality. In order to account for that, we approximate the marginal latent distribution $p(x_t|\mathcal{I})$ as

$$p(x_t|\mathcal{I}) \approx \sum_{v=1}^{V} \pi_v(\mathcal{I}) p(x_{t,v}|\mathcal{I}), \tag{17}$$

where each mixture component $p(x_{t,v}|\mathcal{I})$ is approximated with the BMM algorithm as $p(x_{t,v}|\mathcal{I}) \approx \mathcal{N}(\mu_{t,v}(\mathcal{I}), \Sigma_{t,v}(\mathcal{I}))$. Assuming a GMM at the initial state allows us to efficiently obtain multimodal predictions while being computationally efficient: We obtain multimodal distributions by explicitly marginalizing over the mixture components and we compute each component efficiently by applying the BMM algorithm. Finally, we approximate the term $p(y_t|\mathcal{I})$ as a GMM by another round of moment matching

$$p(y_t|\mathcal{I}) = \int p(y_t|g(x_t), \operatorname{diag}(\Gamma(x_t))) p(x_t|\mathcal{I}) dx_t$$

$$\approx \sum_{v=1}^{V} \pi_v(\mathcal{I}) \mathcal{N}(a_{t,v}(\mathcal{I}), B_{t,v}(\mathcal{I})),$$
(18)

where $a_{t,v}(\mathcal{I})$ and $B_{t,v}(\mathcal{I})$ are the mean and covariance of the v-th mixture component at the t-th time step. These two moments are available as

$$a_{t,v}(\mathcal{I}) = \mathbb{E}[q(x_{t,v})], \qquad B_{t,v}(\mathcal{I}) = \operatorname{Cov}[q(x_{t,v})] + \operatorname{diag}(\mathbb{E}[\Gamma(x_t)]), \qquad (19)$$

which is a direct outcome of the law of the unconscious statistician. We present the pseudocode for computing $p(y_t|\mathcal{I})$ using our method in Algorithm 1.

4.4 Output Moments of Graph Neural Network Layers

In order to use the BMM framework, we need to be able to calculate the first two output moments as well as the expected Jacobian of graph neural network layers. In the following, we derive the analytic expression of the output moments and expected Jacobian for the common graph neural net layers: (i) node-wise affine transformation, and (ii) mean aggregation. Output moments for the ReLU activation are provided in Wu et al. (2019).

Let $x_t^{l,m} \in \mathbb{R}^{D_{x,l}}$ be the node features at layer l of node m at time step t and $x_t^l = \{x_t^{l,m}\}_{m=1}^M$ the set of all node features with $x_t^l \in \mathbb{R}^{MD_{x,l}}$. For the sake of brevity, we have omitted here the index of the mixture

Algorithm 1 Bidimensional Moment Matching in Latent Space

```
Inputs: f(x_t, \mathcal{I})
                                                                                                                                                                                                                                                                                                                 ▶ Mean update
                        L(x_t, \mathcal{I})
                                                                                                                                                                                                                                                                                                       Covariance update
                        g(x_t)
                                                                                                                                                                                                                                                                                                             ▶ Mean emission
                                                                                                                                                                                                                                                                                                   Covariance emission
                                                                                                                                                                                                                                                                                              ▷ Embedding function
Outputs: Approximate marginal distribution p(y_T|\mathcal{I})
\mu_{0,1:V}(\mathcal{I}), \Sigma_{0,1:V}(\mathcal{I}), \pi_v(\mathcal{I}) = h(\mathcal{I})
                                                                                                                                                                                                                                                                            ▷ GMM at initial step, Eq. 8
\mu_{0,1:V}(\mathcal{I}), \Sigma_{0,1:V}(\mathcal{I}), \pi_v(\mathcal{I}) = h(\mathcal{I}) for mixture component v \in \{1, \cdots, V\} do for time step t \in \{0, \cdots, T-1\} do \mu_{t+1,v} \leftarrow \mu_t(\mathcal{I}) + \mathbb{E}[f(x_t, \mathcal{I})] \Sigma_{t+1,v} \leftarrow \Sigma_t(\mathcal{I}) + \operatorname{Cov}[f(x_t, \mathcal{I})] + \operatorname{Cov}[x_t, f(x_t, \mathcal{I}),] + \operatorname{Cov}[x_t, f(x_t, \mathcal{I})]^T + \operatorname{diag}\left(\mathbb{E}[L(x_{t-1}, \mathcal{I})]\right) end for a_{T,v}(\mathcal{I}) \leftarrow \mathbb{E}[g(x_{T,v})] B_{T,v}(\mathcal{I}) \leftarrow \operatorname{Cov}[g(x_{T,v})] + \operatorname{diag}\left(\mathbb{E}[\Gamma(x_T)]\right) end for
                                                                                                                                                                                                                                                                      ▶ Horizontal Moment Matching
                                                                                                                                                                                                                                                                                                                                 ⊳ Eq. 14
                                                                                                                                                                                                                                                                                                                                 ⊳ Eq. 14
                                                                                                                                                                                                                                                                                                                                  ⊳ Eq. 19
                                                                                                                                                                                                                                                                                                                                  ⊳ Eq. 19
end for return \sum_{v=1}^{V} \pi_v(\mathcal{I}) \mathcal{N}(a_{T,v}(\mathcal{I}), B_{T,v}(\mathcal{I}))
```

component v. We denote mean and covariance of a graph with M nodes at layer l and time step t as

$$\mathbb{E}[x_t^l] = \begin{bmatrix} \mathbb{E}[x_t^{l,1}] \\ \vdots \\ \mathbb{E}[x_t^{l,M}] \end{bmatrix}, \qquad \operatorname{Cov}[x_t^l] = \begin{bmatrix} \operatorname{Cov}[x_t^{l,1}, x_t^{l,1}] & \dots & \operatorname{Cov}[x_t^{l,1}, x_t^{l,M}] \\ \vdots & \ddots & \vdots \\ \operatorname{Cov}[x_t^{l,M}, x_t^{l,1}] & \dots & \operatorname{Cov}[x_t^{l,M}, x_t^{l,M}]. \end{bmatrix}, \qquad (20)$$

where $\mathbb{E}[x_t^{l,m}] \in \mathbb{R}^{D_{x,l}}$ and $\operatorname{Cov}[x_t^{l,m}, x_t^{l,m'}] \in \mathbb{R}^{D_{x,l} \times D_{x,l}}$. A typical GNN architecture (see Sec. 2.2) consists of an alternation between the aggregation step, in which information of all neighbors is collected, and the update step, in which the features of the node are updated. For the aggregation step, we derive the output moments for the commonly used mean aggregation operation in Sec. 4.4.3. For the update step, we assume that the neural network is built as a sequence of affine transformations and nonlinearities. The output moments of nonlinear activations are applied independently across agents. As a consequence, their rules do not change when used in the GNN setting and we can use the derivations from Wu et al. (2019). Hence it remains open to derive the output moments for affine transformations, as they are used in GNN context, which we tackle in Sec. 4.4.2. The mean aggregation operation and the node-wise affine operation used in the update step are special cases of a standard affine layer, which we review in Sec. 4.4.1.

4.4.1 Standard Affine Transformation

Suppose, we apply an affine transformation to node m at layer l with state $x_t^{l,m}$

$$x_t^{l+1,m} = W^l x_t^{l,m} + b^l, (21)$$

with weight matrix W^l and bias b^l . The output moments are analytically tractable as

$$\mathbb{E}[\boldsymbol{x}_t^{l+1,m}] = \boldsymbol{W}^l \mathbb{E}[\boldsymbol{x}_t^{l,m}] + \boldsymbol{b}^l, \qquad \qquad \operatorname{Cov}[\boldsymbol{x}_t^{l+1,m}] = \boldsymbol{W}^l \operatorname{Cov}[\boldsymbol{x}_t^{l,m}] (\boldsymbol{W}^l)^T.$$

The expected Jacobian of the affine transformation reads as $J_t^l = W^l$.

4.4.2 Node-Wise Affine Transformation

The node-wise affine transformation applies to each node m at layer l with state $x_t^{l,m}$ the same transformation simultaneously with weight matrix W^l and bias b^l . The node-wise affine transformation can be interpreted as a standard affine transformation acting on the set of all nodes x_t^l as

$$\begin{bmatrix} W^{l} x_{t}^{l,1} + b^{l} \\ W^{l} x_{t}^{l,2} + b^{l} \\ \vdots \\ W^{l} x_{t}^{l,M} + b^{l} \end{bmatrix} = \begin{bmatrix} W^{l} & 0 & \dots & 0 \\ 0 & W^{l} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & W^{l} \end{bmatrix} \begin{bmatrix} x_{t}^{l,1} \\ x_{t}^{l,2} \\ \vdots \\ x_{t}^{l,M} \end{bmatrix} + \begin{bmatrix} b^{l} \\ b^{l} \\ \vdots \\ b^{l} \end{bmatrix} = \underbrace{(I_{M} \otimes W^{l})}_{\hat{W}^{l}} x_{t}^{l} + \underbrace{(\mathbf{1}_{M} \otimes b^{l})}_{\hat{b}^{l}}, \tag{22}$$

where I_M is the identity matrix with shape $M \times M$, $\mathbf{1}_M$ is a vector of ones with shape $M \times 1$, and \otimes is the Kronecker product. The output moments of node-wise affine transformation are

$$\mathbb{E}[\boldsymbol{x}_t^{l+1}] = \hat{W}^l \mathbb{E}[\boldsymbol{x}_t^l] + \hat{b}^l, \qquad \operatorname{Cov}[\boldsymbol{x}_t^{l+1}] = \hat{W}^l \operatorname{Cov}[\boldsymbol{x}_t^l] (\hat{W}^l)^T.$$

Similarly as for the standard affine transformation, the expected Jacobian of node-wise affine transformation is analytically available as $J_t^l = \hat{W}^l$.

4.4.3 Mean Aggregation

A commonly used aggregation operation is the mean aggregator, which calculates the message x_t^{l,\mathcal{N}_m} to node m at time step t at layer l as

$$x_t^{l,\mathcal{N}_m} = \frac{1}{|\mathcal{N}_m|} \sum_{m' \in \mathcal{N}_m} x_t^{l,m'}. \tag{23}$$

Let $x_t^{l,\mathcal{N}}$ be the set of all messages, i.e. $x_t^{l,\mathcal{N}} = \{x_t^{l,\mathcal{N}_m}\}_{m=1}^M$. The mean aggregation can be equivalently written as a linear transformation

$$x_t^{l,\mathcal{N}} = \underbrace{(A \otimes I_{D_{x,l}})}_{\hat{A}} x_t^l. \tag{24}$$

Above $A \in \mathbb{R}^{M \times M}$ denotes the row normalized adjacency matrix, which summarizes the edge information \mathcal{E} in matrix format and $I_{D_{x,l}}$ denotes the identity matrix with dimensionality $D_{x,l} \times D_{x,l}$. The Kronecker product expands the adjacency matrix accordingly to the $D_{x,l}$ -dimensional node features. Hence, the mean aggregation corresponds to a linear transformation with a weight matrix consisting of $M \times M$ blocks, where each block is a diagonal matrix of shape $D_x \times D_x$. Its moments are analytically available as

$$\mathbb{E}[x_t^{l,\mathcal{N}}] = \hat{A}\mathbb{E}[x_t^l], \qquad \operatorname{Cov}[x_t^{l,\mathcal{N}}] = \hat{A}\operatorname{Cov}[x_t^l]\hat{A}^T. \tag{25}$$

The expected Jacobian is available as $J_t^l = \hat{A}$.

4.5 Sparse Covariance Approximation

For settings with a large number of agents M or with a high-dimensional state $x_t^{l,m}$, the application of the BMM algorithm can become computationally expensive. In the following, we review the computational complexity of the BMM algorithm. We assume that the GNN model consists of a mean aggregation step followed by multiple node-wise affine transformations and nonlinearities which is the same architecture that we employ later on in our experiments. The mean aggregation is done for each of the D_x latent states independently, and we denote the maximum hidden layer width with H.

Computing the nonlinearities is cheap as the operation acts elementwise and their effect on the runtime can be neglected during this analysis. The other two operations (see Sec. 4.4.2 and Sec. 4.4.3) can be described by affine operations for which the weight matrices are heavily structured; the mean aggregation step corresponds to a weight matrix consisting of $M \times M$ diagonal blocks, the node-wise affine transformation to a block-diagonal weight matrix with M blocks of shape $H \times H$. Propagation of the full covariance matrix through a neural network (forward cost) has the computational complexity of $\mathcal{O}(M^2H^3+M^2H^2D_x+M^2HD_x^2+M^3D_x^2)$ where the first terms is due to the cost of the $H \times H$ -dimensional node-wise affine transformations in the hidden layers, the second and third term due to the cost of the $H \times D$ -dimensional node-wise affine transformation after the aggregation operation, and the fourth term due to the aggregation operation.

The computational cost of the expected Jacobian is $\mathcal{O}(MH^3 + MH^2D_x + M^2D_x^2)$ and we give its derivation in the following. Let the expected Jacobian of the aggregation operation be $\mathbb{E}[J_t^1]$ and the product of the expected Jacobians of the subsequent neural net layers $\mathbb{E}[J_t^{net}] = \prod_{l=2}^L \mathbb{E}[J_t^l]$. The expected Jacobian of the neural net layers $\mathbb{E}[J_t^{net}]$ is block-diagonal with M blocks of shape $D_x \times D_x$ and its computation takes $\mathcal{O}(MH^3 + MH^2D_x)$ time. Multiplying $\mathbb{E}[J_t^1]$ with $\mathbb{E}[J_t^{net}]$ results in a fully populated matrix where each entry can be computed by a single dot product, due to the structure of its factors, and its computation contributes with $\mathcal{O}(M^2D_x^2)$ to the runtime.

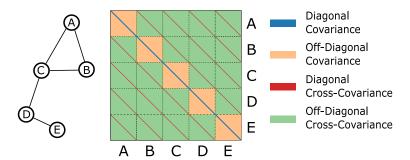


Figure 2: Groupings of the covariance matrix for graph structured data. Consider the example graph with M=5 agents, which is depicted in the left panel. Each agent has dimensionality D_x . In the middle panel, we show the covariance matrix between all agents, which consists of 5×5 blocks, where each block has dimensionality $D_x \times D_x$.

In consequence, the total cost is dominated by the forward cost, $\mathcal{O}(M^2H^3 + M^2H^2D_x + M^2HD_x^2 + M^3D_x^2)$, when using the full covariance matrix. Since the computational cost quickly becomes intractable due to the cubic dependence with respect to the number of agents in the forward pass, we next propose different sparse approximations to the covariance matrix. Independent of the chosen approximation, the cost of the expected Jacobian remains unchanged, as it does not depend on the covariance matrix.

• Full: Model the full covariance matrix. Forward cost: $\mathcal{O}(M^2H^3 + M^2H^2D_x + M^2HD_x^2 + M^3D_x^2)$. Total cost: $\mathcal{O}(M^2H^3 + M^2H^2D_x + M^2HD_x^2 + M^3D_x^2)$.

• Main Diagonal: Keep the diagonal entries in the covariance blocks, which corresponds to the blue line in Fig. 2.

Forward cost: $\mathcal{O}(MH^2 + MHD_x + M^2D_x)$. Total cost: $\mathcal{O}(MH^3 + MH^2D_x + M^2D_x^2)$.

• Main Blocks: Keep the block-diagonal blocks in the covariance matrix, which corresponds to the orange blocks and blue lines in Fig. 2.

Forward cost: $\mathcal{O}(MH^3 + MH^2D_x + MHD_x^2 + M^2D_x^2)$. Total cost: $\mathcal{O}(MH^3 + MH^2D_x + MHD_x^2 + M^2D_x^2)$.

• All Diagonals: Structure the covariance matrix in blocks of shape $M \times M$ and keep the diagonal entries in each block, which corresponds to the blue and red lines in Fig. 2.

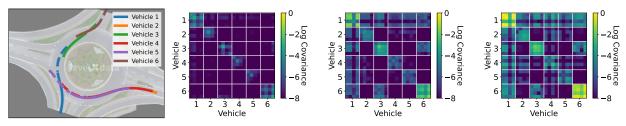
Forward cost: $\mathcal{O}(M^2H^2 + M^2HD_x + M^3D_x)$.

Total cost: $\mathcal{O}(M^2H^2 + M^2HD_x + M^3D_x + MH^3 + MH^2D_x + M^2D_x^2)$.

Note that setting the off-diagonal blocks to zero, as done in Main Blocks and Main Diagonal, corresponds to an independence assumption between the agents and leads to a runtime reduction from $O(M^3)$ to $O(M^2)$. In contrast, assuming a diagonal structure within each block, as performed in Main Diagonal and All Diagonals, corresponds to an independence assumption between the features and leads to runtime reduction from $O(H^3)$ to $O(H^2)$ in the forward pass.

Finally, it is important to note that the covariance matrix does only have the same structure as the graph after the first time step. Agents that are not connected via an edge can still have a non-zero cross-covariance at time step t, provided that they are connected by a path that is at most t steps long. In our applications, this leads to non-sparse covariances after a few time steps, since the number of agents is small compared to the time horizon. We present the covariance matrix at three different time steps for an exemplary scene in Fig. 3. For a short prediction horizon of one second, the covariance matrix has an approximately diagonal

¹For reference, taking a Monte Carlo approach has the computational complexity $\mathcal{O}(SMH^2 + SMHD_x + SM^2D_x)$ where S is the number of Monte Carlo samples.



- (a) Example Scene.
- (b) Covariance at 1 second. (c) Covariance at 3 seconds. (d) Covariance at 5 seconds.

Figure 3: We visualize an example scene from the rounD dataset (Krajewski et al., 2020) and the covariance matrix in the latent space at different time steps for the case of a unimodal initial distribution. In the left plot, dashed lines represent the observed history and solid lines represent the true future trajectories.

shape. As the prediction horizon increases, the covariance matrix becomes more complex and is no longer dominated by its diagonal entries. We remark that the we could further increase the information spread across agents by using an architecture with multiple aggregation steps within the GNN module if required.

5 Experiments

We provide experiments on two challenging autonomous driving datasets. The first experiment (Sec. 5.1) conducts an ablation study, while the second experiment (Sec. 5.2) benchmarks our model against state-of-the-art methods. We use no covariance approximation in the first two experiments, i.e. we use the full covariance matrix. We provide a runtime analysis and a benchmark of our proposed sparse covariance approximations in Sec. 5.3.

We apply the same architecture for the embedding, transition, and emission function for rounD and NGSIM experiments, which we present in App. A.

5.1 rounD

The rounD dataset (Krajewski et al., 2020) consists of vehicle trajectories recorded at different roundabouts in Germany. As the roundabouts involve many interactions among vehicles, we expect the predictive distributions to be multimodal and highly complex. We use the recordings from the roundabout in Neuweiler near Aachen for training and testing purposes. The dataset consists of 13,129 tracked objects recorded at 25 Hz in 22 sessions amounting to a total recording time of 6.6 hours. We remove pedestrians, bicycles, as well as parked vehicles from the dataset as their influence on the vehicle behaviour patterns in the roundabouts is negligible. After dataset curation, we are left with 12,715 tracked objects. We downsample the recordings by a factor of five and construct a dataset consisting of eight seconds long segments with 50 % overlap, resulting in 5405 snippets. We use the first three seconds as the track history, which is part of the context variable \mathcal{I} , and the following five seconds as the prediction horizon. The first 18 recording sessions, corresponding to 4,314 snippets, are used for training and validation. The final four recording sessions, corresponding to 1,091 snippets, are used for testing. We build the connection graph by connecting vehicles with an Euclidean distance less than 30 meters.

5.1.1 Baselines

We compare our method with multiple baselines. For each baseline, we remove one key assumption of our model. We cite papers that employ similar ideas as appropriate. We did not reimplement these works but performed an ablation study in which we replaced specific components of our model in the interest of a fair comparison.

(i) Monte Carlo (MC): Our model with Monte Carlo based training. One forward pass through our model amounts approximately to the computational cost of 12 Monte Carlo simulations. For a fair comparison, we

use during training 16 particles, which is more costly than training with our proposed moment propagation algorithm, and test with 100 particles.

- (ii) Multimodality: We increase the number of components in the GMM prior in order to test the effect of multimodality.
- (iii) Linearity (e.g. Li et al. (2020)): We remove all non-linearities from the latent dynamics. Note that we keep the non-linear emission model in order to map the dynamics into a latent space in which the system can be linearly approximated.
- (iv) No Interactions (e.g. Krishnan et al. (2017)): Our model with a diagonal adjacency matrix, i.e. we remove all edges from the graphs. This model neglects interactions between traffic participants.
- (v) Non Recurrent GNN (e.g. Casas et al. (2020); Herman et al. (2021)): This architecture receives the context variable \mathcal{I} , performs one round of message passing, and subsequently outputs one normal distribution for each of the next five seconds without using a recurrent architecture.
- (vi) No Latent Noise (e.g. Sanchez-Gonzalez et al. (2020)): We remove the noise from the latent dynamics [Eq. (2)] while keeping the emission model unchanged. The uncertainty can no longer be propagated forward in time as the emission model acts independently for each time point.

5.1.2 Results

We provide benchmark results of all methods in Tab. 1. First, we compare our deterministic training and testing scheme against its Monte Carlo alternative. Though Monte Carlo based training is more costly than training with BMM, the Monte Carlo results are significantly outperformed by our method. Our results indicate that our deterministic approach leads to more effective approximations compared to Monte Carlo sampling despite the approximation error we obtain by our deterministic moment matching scheme. One potential explanation for our finding is that the Monte Carlo approaches suffer under a high variance since the latent space for multi-agent space grows linearly with the number of agents.

Next, we study if our method can capture multimodality by increasing the number of components in the GMM prior. It is worth noting that GMMs can approximate arbitrarily complex distributions when the number of components is chosen high enough. In our experiments, we observe that an increase of components significantly decreases the *Root Mean Squared Error* (RMSE) and NLL (negative log likelihood) making the GMM prior a vital ingredient of our method and suggesting that the true predictive distribution is highly multi-modal. If the number of components is chosen too small, our model adjusts its uncertainty predictions accordingly. For example, we observe in Fig. 1 the uncertainty of the orange agent to increase as the vehicle is close to the exit of the roundabout. The high predictive uncertainty can be explained by two potential future outcomes: the orange vehicle can leave the roundabout or stay inside. Consequently, our model learns to compensate if the number of components is picked too low which in turn allows us to trade accuracy for computational runtime. Last but not least, when using tailored implementations, one can easily scale up to a larger number of components since their computations can be parallelized without any hurdles.

We then compare our model to a simpler alternative in which the dynamics are assumed to be linear which allows calculating the moments of the transition model exactly and in closed form (Särkkä & Solin, 2019). In contrast, our approach, GDSSM, approximates the non-linear dynamics in a local linear way by using deterministic moment matching results. We find that our approach achieves lower RMSE and NLL, which can most likely be attributed to the higher modeling flexibility of our proposed model class.

Our ablation study further shows discarding latent noise, modeling the dynamics in a non-recurrent manner or removing interactions between traffic participants results in higher RMSE and NLL.

5.2 NGSIM

The Next Generation Simulation (NGSIM) dataset (Halkias & Colyar, 2007) consists of vehicle trajectories recorded at 10 Hz at two different highways, US-101 and I-80, in the United States. The dataset is commonly used for benchmarking traffic forecasting methods and allows us to compare our method against prior art.

Table 1: RounD results. We provide RMSE and NLL (mean \pm standard error over 10 runs). For the case of predictors with multiple modes, we calculate the RMSE for each mode and report the lowest.

		Non Recurrent GNN	GDSSM 1 Mode Det.	GDSSM 1 Mode Det.	GDSSM 1 Mode Det.	GDSSM 1 Mode MC	GDSSM 1 Mode Det.	GDSSM 2 Modes Det.	GDSSM 3 Modes Det.	GDSSM 4 Modes Det.
		1 Mode	No Lat. Noise	Linear	No Interaction	MC	Det.	Det.	Det.	Det.
RMSE	1 s 2 s 3 s 4 s 5 s	$ \begin{vmatrix} 0.72 \pm 0.02 \\ 1.90 \pm 0.02 \\ 3.54 \pm 0.03 \\ 5.26 \pm 0.04 \\ 7.20 \pm 0.05 \end{vmatrix} $	$\begin{array}{c} 1.22 \pm 0.08 \\ 2.33 \pm 0.08 \\ 3.88 \pm 0.07 \\ 5.58 \pm 0.08 \\ 7.65 \pm 0.10 \end{array}$	$ \begin{vmatrix} 0.88 \pm 0.02 \\ 2.12 \pm 0.03 \\ 3.88 \pm 0.05 \\ 6.13 \pm 0.09 \\ 8.83 \pm 0.13 \end{vmatrix} $	$\begin{array}{c} 0.91 \pm 0.03 \\ 2.10 \pm 0.04 \\ 3.69 \pm 0.06 \\ 5.39 \pm 0.13 \\ 7.56 \pm 0.23 \end{array}$	$\begin{array}{c} 0.90 \pm 0.02 \\ 2.09 \pm 0.02 \\ 3.60 \pm 0.04 \\ 5.37 \pm 0.05 \\ 7.65 \pm 0.06 \end{array}$	$ \begin{vmatrix} 0.79 \pm 0.02 \\ 1.87 \pm 0.02 \\ 3.36 \pm 0.03 \\ 5.08 \pm 0.04 \\ 7.24 \pm 0.05 \end{vmatrix} $	$ \begin{vmatrix} 0.75 \pm 0.04 \\ 1.85 \pm 0.06 \\ 3.46 \pm 0.11 \\ 5.07 \pm 0.13 \\ 6.29 \pm 0.23 \end{vmatrix} $	$ \begin{vmatrix} 0.76 \pm 0.03 \\ 1.83 \pm 0.07 \\ 3.05 \pm 0.16 \\ 4.55 \pm 0.28 \\ 5.95 \pm 0.47 \end{vmatrix} $	$ \begin{vmatrix} 0.69 \pm 0.02 \\ 1.68 \pm 0.05 \\ 2.75 \pm 0.12 \\ 4.04 \pm 0.45 \\ 4.43 \pm 0.53 \end{vmatrix} $
NLL	1 s 2 s 3 s 4 s 5 s	$ \begin{vmatrix} 1.90 \pm 0.01 \\ 3.25 \pm 0.02 \\ 4.40 \pm 0.02 \\ 5.13 \pm 0.02 \\ 5.71 \pm 0.02 \end{vmatrix} $	$\begin{array}{c} 2.96 \pm 0.08 \\ 3.85 \pm 0.10 \\ 5.05 \pm 0.13 \\ 6.17 \pm 0.16 \\ 7.24 \pm 0.20 \end{array}$	$\begin{array}{c} 1.95 \pm 0.08 \\ 3.93 \pm 0.13 \\ 5.11 \pm 0.19 \\ 6.01 \pm 0.22 \\ 6.80 \pm 0.22 \end{array}$	$ \begin{array}{c} 1.67 \pm 0.07 \\ 3.37 \pm 0.08 \\ 4.34 \pm 0.08 \\ 4.93 \pm 0.09 \\ 5.51 \pm 0.10 \end{array} $	$\begin{array}{c} 2.82 \pm 0.03 \\ 4.11 \pm 0.02 \\ 4.67 \pm 0.09 \\ 5.01 \pm 0.07 \\ 5.41 \pm 0.05 \end{array}$	$ \begin{vmatrix} 1.48 \pm 0.05 \\ 2.91 \pm 0.03 \\ 3.87 \pm 0.02 \\ 4.46 \pm 0.03 \\ 5.05 \pm 0.04 \end{vmatrix} $	$\begin{array}{c} 1.34 \pm 0.08 \\ 2.93 \pm 0.07 \\ 4.01 \pm 0.07 \\ 4.52 \pm 0.11 \\ 4.82 \pm 0.24 \end{array}$	$ \begin{vmatrix} 1.36 \pm 0.04 \\ 2.93 \pm 0.06 \\ 3.77 \pm 0.10 \\ 4.21 \pm 0.15 \\ 4.59 \pm 0.15 \end{vmatrix} $	$ \begin{vmatrix} 1.21 \pm 0.05 \\ 2.79 \pm 0.09 \\ 3.83 \pm 0.08 \\ 4.34 \pm 0.18 \\ 4.27 \pm 0.23 \end{vmatrix} $

We adopt the experimental setup of Deo & Trivedi (2018) and use both highway scenarios. We split the scenarios into three 15 minute long time spans resembling mild, moderate, and congested traffic conditions and downsample each trajectory by a factor of two. The test set consists of a fourth of all trajectories randomly sampled from both locations. As in the rounD experiment, we split each trajectory into eight seconds long segments, where the first three seconds are used as the track history and the following five seconds as the prediction horizon.

Similarly as in Diehl et al. (2019); Lenz et al. (2017); Wheeler & Kochenderfer (2016), we introduce a connection graph based on the lane position of each vehicle. Each vehicle has at most six connections to other vehicles, which are the nearest vehicles in front/behind on the same/left/right lane. The vehicles on the outermost lanes have a maximum of four neighbours, as there are no neighbours to the left or right.

5.2.1 Baselines

There exists a large body of prior work (Su et al., 2020; Jeon et al., 2020; Mo et al., 2021; Diehl et al., 2019; Wheeler & Kochenderfer, 2016; Lenz et al., 2017), which applies GNNs for traffic forecasting on the NGSIM dataset and provide quantification of predictive uncertainty. We compare our approach with the following methods:

- (i) Constant Velocity (Mercat et al., 2019): This method uses a linear state-space model together with a Kalman filter in order to make predictions. It does not take interactions between agents into account.
- (ii) Convolutional Social (CS)-LSTM (Deo & Trivedi, 2018): Interactions between vehicles are modeled by introducing a grid and applying a convolutional layer on top. Dynamics are modeled by a deterministic LSTM, which predicts the mean and the variance of a normal distribution at each time step.
- (iii) Multiple Futures Prediction (MFP) (Tang & Salakhutdinov, 2019): A recurrent model with deterministic transition dynamics. Stochasticity is introduced via the initial state and noisy observations that are fed back into the dynamical model. Interactions are modeled by an attention module.

To the best of our knowledge, no prior work uses a stochastic GNN and provides uncertainty quantification results on this dataset. We therefore introduce as additional baseline our method using Monte Carlo simulations.

5.2.2 Results

We provide benchmark results of our proposed model in Tab. 2. Similar to the experiments on the rounD dataset in Sec. 5.1, we observe that (i) deterministic training and testing is more efficient than its Monte Carlo based alternative and (ii) increasing the number of modes improves the performance.

Next, we compare our method, using one component in the GMM prior only, with all other unimodal prediction methods. We can observe that our approach outperforms its competitors in terms of negative log-likelihood, while MFP achieves a smaller test RMSE.

We subsequently increase the number of modes for our model and for MFP. For our method, an increase in modes leads to a significant improvement in both metrics, while MFP primarily improves in NLL. In consequence, our method achieves the smallest RMSE compared to all other methods over all time horizons. In terms of NLL, our method achieves superior results when it comes to long-term predictions (3s, 4s, 5s), while MFP performs better for short-term predictions (1s, 2s). Accurate long-term prediction of the agents in a driving scene is crucial for high-level autonomous driving, as an accurate environment model is a prerequisite for precise planning of driving controls. For instance, advanced driver-assistance systems usually use time horizons between 3s and 5s for driver warnings and emergency brakes, while autonomous cars aim for a time horizon for 5s or longer in order to ensure safe and comfortable rides (Philipp & Goehring, 2019).

Table 2: NGSIM results. We provide average and standard error over 10 runs. For the case of predictors with multiple modes we calculate the RMSE for each mode and report the lowest.

		Constant Velocity	CS-LSTM	MFP 1 Mode	MFP 4 Modes	GDSSM 1 Mode MC	GDSSM 1 Mode Det.	GDSSM 2 Modes Det.	GDSSM 3 Modes Det.	GDSSM 4 Modes Det.
	$1\mathrm{s}$	0.75	0.61	0.54	0.54	0.56 ± 0.00	0.53 ± 0.01	0.46 ± 0.01	0.35 ± 0.01	0.32 ± 0.01
闰	$2\mathrm{s}$	1.81	1.27	1.16	1.17	1.27 ± 0.02	1.18 ± 0.01	1.05 ± 0.01	0.80 ± 0.01	0.71 ± 0.01
RMSE	$3\mathrm{s}$	3.16	2.09	1.90	1.91	2.11 ± 0.03	1.98 ± 0.02	1.73 ± 0.02	1.33 ± 0.02	1.18 ± 0.03
	$4\mathrm{s}$	4.80	3.10	2.78	2.75	3.16 ± 0.04	2.99 ± 0.03	2.60 ± 0.04	2.02 ± 0.04	1.75 ± 0.04
	$5\mathrm{s}$	6.70	4.37	3.83	3.78	4.47 ± 0.05	4.29 ± 0.04	3.69 ± 0.06	2.88 ± 0.05	2.46 ± 0.05
	$1\mathrm{s}$	0.80	0.58	0.73	-0.65	0.64 ± 0.04	0.19 ± 0.02	-0.12 ± 0.02	-0.15 ± 0.03	-0.16 ± 0.03
,	$2\mathrm{s}$	2.30	2.14	2.33	1.19	2.10 ± 0.10	1.61 ± 0.02	1.37 ± 0.02	1.35 ± 0.02	1.32 ± 0.02
NLI	$3\mathrm{s}$	3.21	3.03	3.17	2.28	2.92 ± 0.11	2.42 ± 0.02	2.23 ± 0.02	2.23 ± 0.02	2.19 ± 0.02
	$4\mathrm{s}$	3.89	3.68	3.77	3.06	3.54 ± 0.10	3.02 ± 0.02	2.88 ± 0.02	2.88 ± 0.02	2.82 ± 0.02
	$5\mathrm{s}$	4.44	4.22	4.26	3.69	4.04 ± 0.10	3.50 ± 0.02	3.42 ± 0.02	3.41 ± 0.02	3.36 ± 0.02

5.3 Covariance Approximations

We provide a detailed runtime analysis for different covariance approximations in Sec. 5.3.1 and study their impact on the predictive performance in Sec. 5.3.2.

5.3.1 Runtime

We visualize the runtime of different covariance approximations, as well as the runtime of the Monte Carlo alternative in Fig. 4 as a function of input dimensionality and number of agents. We use the same NSDE architecture as in our experiments on the rounD and NGSIM dataset. For the Monte Carlo alternative, we visualize the runtime for 16 particles, as we use the same number of particles for training in Sec. 5.1 and 5.2.

We first confirm that propagation the full covariance matrix is more costly than any of the proposed approximations. In fact, our sparse covariance approximations can reduce the runtime up to a factor of 100 for systems with a large number of agents and a high input dimensionality. As we derived in Sec. 4.5, when using the BMM algorithm with a full covariance matrix or the all diagonals approximation, the computational cost shows a cubic dependence on the number of agents. In contrast, the main diagonal approximation, the main blocks approximation, as well as MC based predictions have a quadratic dependence on the number of agents. This makes these two approximations an attractive alternative to MC based predictions, when systems with a high number of agents need to be modeled with a limited computational budget.

For systems with a moderate input dimensionality (≈ 8) and number of agents (≈ 16), all of our proposed approximations require only up to 5 ms in order to compute the distribution at the next time point, which corresponds to the cost of 5-10 Monte Carlo simulations.

5.3.2 Benchmark

Next, we study the effect of different covariance approximations on the performance. We report the results for the case of a unimodal GDSSM. The results are depicted in Tab. 3. We report here our main findings.

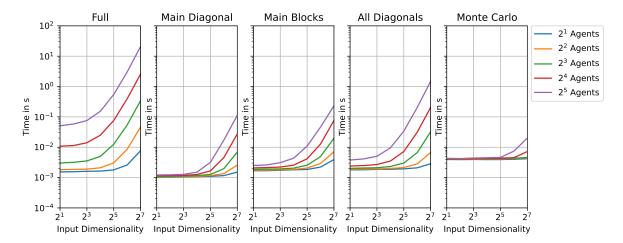


Figure 4: Wallclock time for output moment calculation for the latent dynamics using GNNs with three hidden layers of size 24 for different covariance approximations (from left to right). For each approximation, we plot the runtime as a function of the input dimensionality D_x and number of agents M. The GNNs are initialized at random and we report the average runtime over 100 repetitions.

First, modeling the full covariance matrix results in the best performance in terms of lowest RMSE and NLL. Second, the all diagonals covariance approximation performs the best among the covariance approximations. It achieves comparable MSE as the full solution, but falls slightly behind in NLL when the system is highly interactive (see rounD dataset). In this setting, it outperforms the other two sparse approximations with respect to NLL. This behavior might be explained by the assumptions made in the covariance approximation: it is the only approximation that allows for correlations between agents as its structure only neglects dependencies between latent features.

Third, the differences in performances between the full solution and the different approximations with respect to MSE lie between one and two standard errors. Modeling the main diagonal only can thus be sufficient for applications with low computational resources and a high demand in accuracy by accepting a slight loss in calibration. For these applications, the runtime can be reduced from $O(M^3)$ to $O(M^2)$.

Table 3: Test performance for different covariance approximations on the rounD and NGSIM dataset for the unimodal case. We provide average and standard error over 10 runs.

			rou	nD		NGSIM				
		Full	Main Diagonal	Main Blocks	All Diagonals	Full	Main Diagonal	Main Blocks	All Diagonals	
RMSE	$1\mathrm{s}$	0.79 ± 0.02	0.82 ± 0.02	0.83 ± 0.04	0.78 ± 0.02	0.53 ± 0.01	0.54 ± 0.01	0.55 ± 0.01	0.53 ± 0.01	
	$2\mathrm{s}$	1.87 ± 0.02	1.88 ± 0.02	1.89 ± 0.05	1.87 ± 0.02	1.18 ± 0.01	1.18 ± 0.02	1.19 ± 0.02	1.19 ± 0.02	
	$3\mathrm{s}$	3.36 ± 0.03	3.40 ± 0.02	3.38 ± 0.06	3.34 ± 0.02	1.98 ± 0.02	1.99 ± 0.03	2.05 ± 0.03	1.99 ± 0.02	
	$4\mathrm{s}$	5.08 ± 0.04	5.07 ± 0.04	5.09 ± 0.07	5.05 ± 0.03	2.99 ± 0.03	2.98 ± 0.04	3.07 ± 0.04	2.97 ± 0.03	
	$5\mathrm{s}$	7.24 ± 0.05	7.25 ± 0.06	7.30 ± 0.08	7.25 ± 0.05	4.29 ± 0.04	4.34 ± 0.05	4.54 ± 0.06	4.32 ± 0.04	
NEL	$1\mathrm{s}$	1.48 ± 0.05	1.77 ± 0.06	1.79 ± 0.05	1.69 ± 0.03	0.19 ± 0.02	0.24 ± 0.04	0.22 ± 0.03	0.18 ± 0.03	
	$2\mathrm{s}$	2.91 ± 0.03	3.34 ± 0.04	3.35 ± 0.06	3.25 ± 0.02	1.61 ± 0.02	1.63 ± 0.03	1.64 ± 0.03	1.59 ± 0.03	
	$3\mathrm{s}$	3.87 ± 0.02	4.27 ± 0.03	4.28 ± 0.05	4.18 ± 0.02	2.42 ± 0.02	2.44 ± 0.02	2.46 ± 0.03	2.43 ± 0.02	
	$4\mathrm{s}$	4.46 ± 0.03	5.00 ± 0.02	5.01 ± 0.05	4.92 ± 0.02	3.02 ± 0.02	3.04 ± 0.02	3.06 ± 0.04	3.01 ± 0.02	
	$5\mathrm{s}$	5.05 ± 0.04	5.69 ± 0.02	5.68 ± 0.06	5.64 ± 0.05	3.50 ± 0.02	3.59 ± 0.02	3.62 ± 0.03	3.57 ± 0.02	

6 Conclusion

In this work, we have proposed GDSSMs in which the latent dynamics of the agents are coupled via GNNs in order to capture interactions among multiple agents. We derived moment matching rules for GNN layers that allow for deterministic inference and introduced a GMM prior over the initial latent states in order to allow for multimodal predictions. Both together lead to an efficient and stable algorithm that is able

to produce complex and nonlinear predictive distributions. We experimentally validated that our novel method achieves accurate and well-calibrated predictions on two challenging autonomous driving datasets. Finally, we proposed sparse approximations to the covariance matrix considering the computational limits of real-world vehicle control units. Depending on the required calibration, our approximations can lead to a significant reduction of the runtime without impeding accuracy.

In future work, we seek to increase the robustness of our proposed model towards novel and unseen traffic scenarios. One way could be to incorporate epistemic uncertainty into our model formulation by placing a prior over the weights of the GNN. To achieve this, we could combine our model with recent advances in variational inference in order to find an approximation to the intractable weight posterior. Another research direction of interest is modeling of irregular and partially observed dynamical systems. Prior work uses continuous time encoder networks as well as a continuous time transition model in latent space (Rubanova et al., 2019; Brouwer et al., 2019). Following this vein of work, an extension of our models towards continuous time networks seems a promising direction for modeling interactive systems.

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A Network Architectures

Below we present the neural network architectures for rounD and NGSIM experiments. The embedding function receives the history of all M agents. This history is 3 seconds long with a time step of 0.2 seconds and consists of two dimensional coordinates. After flattening, the input is a vector of size 30. The output of the embedding function is a GMM with V mixture components. We use the mean aggregator in the embedding, mean and variance update function. The mean aggregator calculates the message to agent m accordingly to Eq. 23. After the message $x_t^{\mathcal{N}_m}$ is calculated, it is concatenated with the state x_t^m . Mean and variance update functions are neural networks, which conduct at each prediction step one round of message passing and then calculate the output. Our emissions model uses a neural network for the mean function $g(x_t)$ and a constant vector for $Q(x_t)$. The emission model maps the state of each agent back to the observed space and does not depend on interactions.

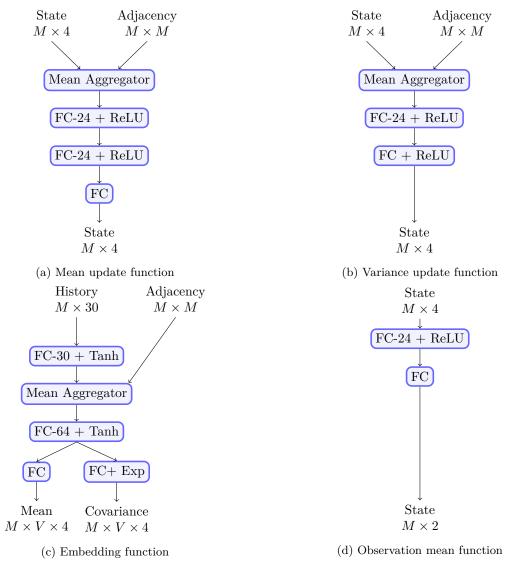


Figure 5: Architectures.