
Unifying (Federated) (Private) High-Dimensional Bandits via ADMM

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Abstract

1 We study all possible variants of the high dimensional stochastic linear contextual
2 bandit problem in federated and private settings. We propose a unifying algorithm
3 design and analysis framework built on ADMM. Our method achieves existing
4 state-of-the art guarantees in either setting for the central model. For the federated
5 model, our results are entirely new and near-optimal in either setting. We also
6 establish a novel lower bound on privacy-utility tradeoff for the federated model
7 in the private setting and demonstrate on suitable numerical experiments for all
8 problem variants.

9 1 Introduction

10 We study a high-dimensional stochastic contextual linear bandit problem in central and federated
11 learning settings under privacy constraints (Shukla, 2024; Chakraborty et al., 2024). At every epoch,
12 the decision maker is given a set of stochastically generated exogeneous contexts and chooses from
13 a finite set of actions $[K]$ to obtain the highest reward. Since the decision maker does not know
14 the true underlying reward, it creates an estimator to select a decision. This induces an exploration-
15 exploitation trade-off wherein the decision-maker faces the dilemma of exploring arms not played
16 before and exploiting the information accumulated so far. The key underlying assumption in this
17 framework is that the unknown parameter vector is s^* -sparse. This paper is motivated by the need
18 for a unified algorithm and analysis framework for establishing utility and privacy guarantees for
19 this problem under different learning models. From the perspective of this paper, existing work
20 in high-dimensional bandits can be seen in central or federated settings with or without privacy
21 constraints. We motivate the need for such a framework in the in Example 1.1 this is a very common
22 learning setting.

23 **Example 1.1.** *A crucial step in assuring the drug safety of oligonucleotide drugs requires learning*
24 *the relevant thermodynamics from large-scale data distributed across different organizations (Tavara*
25 *et al., 2021). Preserving data privacy in this distributed setup requires limited and private communi-*
26 *cation between local nodes. Further, learning the safety curve online requires learning in an online*
27 *fashion. Techniques proposed in this paper can be used to solve this problem of online, private, and*
28 *federated learning with high-dimensional but sparse data.*

29 In the central setting with privacy constraints, the regret analysis follows the by-now standard
30 framework: (i) constructing an estimator for the unknown parameter using LASSO or thresholding-
31 based techniques, (ii) analysing per-step regret using underlying problem structure (such as margin
32 condition, compatibilty or other similar conditions). In this paper, we eschew this line of arguments
33 and propose a generic algorithmic design and analysis framework for this class of problems. In either
34 settings, we view the statistical problem of constructing the LASSO-based estimator through an
35 optimizer’s lens. By exploiting the architecture of ADMM-based optimizers (Boyd et al., 2011),
36 our proposed method can achieve state-of-the-art guarantees under all possible combinations of the

Table 1: Contributions to existing work. New results are in blue. $\rho = \epsilon^{-2} \log(1/\delta)$ for (ϵ, δ) -DP.

		Central (Theorems 2, 3)	Federated (Theorems 2, 4)
Non-Private	Lower bound	$\Omega(\sqrt{s^* T \log(d/s^*)})$	$\Omega(\sqrt{s^* M T \log(d/s^*)})$
	Upper bound (existing)	$\mathcal{O}(\sqrt{s^* T \log(d/s^*)})$	–
	Upper bound (ours)	$\mathcal{O}(s^* \sqrt{T \log(d/s^*)})$	$\mathcal{O}(s^* \sqrt{M T \log(d/s^*)})$
Private $((\epsilon, \delta)$ -DP)	Lower bound	$\Omega(\log(d/s^*) \sqrt{s^* \rho})$	$\Omega(\log(d/s^*) \sqrt{s^* \rho})$
	Upper bound (existing)	$\mathcal{O}(\sqrt{\rho}(s^*)^{1.5} \log^{1.5}(d/s^*))$	–
	Upper bound (ours)	$\mathcal{O}(\sqrt{\rho}(s^*)^{1.5} \log^2(d/s^*))$	$\mathcal{O}(\sqrt{\rho}(s^*)^{1.5} \log^2(d/s^*))$

learning models and privacy settings (see Table 1). Detailed literature is reviewed in Appendix A. In high-dimensional bandits in the central model, closest to our work are (Chakraborty et al., 2024; Shukla, 2024) that consider the high-dimensional bandit problem in the central model with privacy and propose different thresholding-based methods to solve it. For the federated setup the non-private setting was considered by (Wang et al., 2023). To the best of our knowledge, there is no known work on private high-dimensional federated contextual bandits. This motivates the question:

Is it possible to propose a generic algorithm design and analysis template for high-dimensional bandits that can operate both in central and federated settings, and with and without differential privacy constraints?

Our Contributions: We address these questions affirmatively and contribute to all the aforementioned strands of literature. Our contributions are summarised in Table 1 and detailed as follows:

1. We propose a general framework for designing algorithms for peeled LASSO (Section 3) and high-dimensional bandit problems (Section 4) applicable to several variants of the problem that have been considered independently until now. These include centralized/federated and private/non-private versions thereof. Our framework works for all combinations of these variants by tuning a few hyper-parameters.
2. In Algorithms 1, we propose an admm-based algorithm for the PeeledLASSO applicable for either communication model and establish its estimation error in Theorem 1. These are the first known recovery guarantees for the online private LASSO in private and federated setting. Algorithm 3 extends this to an algorithm for high-dimensional bandits in either setting using forgetting and the doubling trick. Privacy guarantees for these algorithms are established in Theorem 2 and utility guarantees in Theorem 3 and 4. Our utility proofs non-trivially combine iteration-based analysis of ADMM and peeling-based privacy arguments to accommodate bandit-feedback under both communication models.
3. In Theorem 5, we derive a problem independent regret lower bound for the federated model in both private and non-private setting. Our lower bound demonstrates a phase transition between the hardness of the problem depending on the privacy level, intrinsic dimension and sparsity. In the low-dimensional case, a similar phenomenon was shown by Azize and Basu (2022).

2 Model and Preliminaries

Central Model: We consider a linear contextual bandit problem with K arms, where the (unknown) underlying vector $\theta \in \mathbb{R}^d$ that parameterizes rewards is high-dimensional and sparse – often called the LASSO bandit setting (Bastani and Bayati, 2020). Central and federated models correspond to different modes of communication between central server and agents. We consider a time horizon T , where for each time $t \in [T]$, the algorithm is given an (exogenous) context vector $\mathcal{X}_t = \{X_{t,k}\}_{k \in [K]}$, where each $X_{t,k} \in \mathbb{R}^d$. The elements of the set \mathcal{X}_t are drawn i.i.d. from an unknown distribution \mathcal{D} . After observing \mathcal{X}_t , the algorithm selects an arm $k_t \in [K]$ and observes a random reward r_t given by $r_t = \langle X_{t,k_t}, \theta \rangle + \eta_t$, where η_t is zero-mean bounded noise with variance σ^2 . Let \mathcal{H}_t denote the tuple of random variables generated by the past contexts, arm pulls, and observed rewards and \mathcal{F}_t denote the corresponding filtration. We assume that the parameter vector $\theta \in \mathbb{R}^d$ is s -sparse, i.e., its support $S = \{i \in [d] : \theta_i \neq 0\}$ has cardinality s , which is known to the decision maker. Given the history of observations $\mathcal{H}_t = \{X_{k_s, s}, r_s\}_{s=1}^t$, let \mathcal{F}_t denote the natural filtration associated with \mathcal{H}_t and let Π denote the set of all \mathcal{F}_t -measurable policies. In the central setting,

our goal is to design a policy $\pi \in \Pi$ that minimizes the cumulative expected regret, defined as:
 $R^\pi(T) = \sum_{t=1}^T \mathbb{E}[\max_{k \in [K]} \langle X_{t,k} - X_{t,k_t}, \theta \rangle].$

Federated Model: In the federated setup, $[M]$ is the set of clients, $[K]$ is the set of arms, \mathcal{X} is the set of contexts. Each client is a K -armed bandit problem with a common parameter being shared across clients. At each time slot t , each client i observes a exogeneous contexts $X_{i,t} \in \mathcal{X}$, pulls arm $k_{i,t} \in [K]$ and receives a reward $r_{i,t} = X_{i,t}^\top \theta + \eta_{i,t}$. We assume there exists a central server in the system, and similar to FL, the clients can communicate with the server periodically with zero latency. Specifically, the clients can send “local model updates” to the central server, which then aggregates and broadcasts the updated “global model” to the clients. We also assume that clients and server are fully synchronized. The class of policies considered here are restricted to collinearly correlated policies Huang et al. (2021) (see Definition 6 in Appendix H). Intuitively, for two clients that are not collinear, their local observations on any arm cannot be utilized to improve each other’s knowledge of their own local models. As a result, they should not affect each other’s decision-making process. For a collinearly correlated policy, the regret in this setting is defined as:
 $R^\pi(T) = \sum_{s=1}^T \sum_{i \in [M]} \left(X_{i,k_t^*}^\top \theta - X_{i,k_t}^\top \theta \right).$

Privacy: Differential privacy Dwork et al. (2010) is the most prominent notion of privacy. The contextual bandit problem involves two sets of variables that any agent must private to the other participating agents – the available decision sets $\{X_{k,s}\}_{k \in [K], s \in [T]}$ and observed rewards $\{r_s\}_{s \in [T]}$. In our problem, we are concerned with preserving privacy of continual observations Dwork et al. (2010) under different communication protocols, center and federated and bandit feedback. We instead use *joint differential privacy* (JDP), first introduced by Kearns et al. (2014) in the context of algorithmic mechanism design, and later extended to the online bandit setting by Shariff and Sheffet (2018). This slight relaxation of differential privacy allows the t -th component of the output (i.e., k_t) to depend arbitrarily on the t -th component of the input (i.e., \mathcal{X}_t), while ensuring differential privacy with respect to the joint distribution of all other components of the output.

Definition 1 (Joint Differential Privacy (Kearns et al., 2014; Shariff and Sheffet, 2018)). *A streaming algorithm $\mathcal{A} : \mathcal{H}^T \rightarrow [K]^T$ is said to be (ϵ, δ) -jointly differential private, for any t -neighboring streams $\mathcal{S}, \mathcal{S}'$ and any $\mathcal{T} \subseteq [K]^{T-1}$,*

$$\Pr[\mathcal{A}(\mathcal{S})_{\neq t} \in \mathcal{T}] \leq \exp(\epsilon) \Pr[\mathcal{A}(\mathcal{S}')_{\neq t} \in \mathcal{T}] + \delta,$$

where $\mathcal{A}(\mathcal{S})_{\neq t}$ denotes all portions of the algorithm’s output except at time t .

The adversary model assumed here is to prevent any two colluding agents j and k to obtain non-private information about any specific element in agent i ’s history. Therefore, the context set $\mathcal{X}_{i,t}$ and outcome $r_{i,t}$ are sensitive variables that the user trusts only with the agent i . Hence, we wish to keep $\mathcal{X}_{i,t} \in [T]$ private. However, the agent only stores the chosen actions $X_{i,t}$ (and not all of $\mathcal{X}_{i,t}$), and hence making our technique differentially private with respect to $((x_{i,t}, r_{i,t}), t \in [T])$ will suffice.

Definition 2 $((\epsilon, \delta)$ -Fed-JDP (Dubey and Pentland, 2020)). *In a federated learning setting with $M \geq 2$ agents, a randomized multi-agent contextual bandit algorithm $A = (A_i)$, $M_i = 1$ is (ϵ, δ, M) -federated differentially private under continual multi-agent observation if for any i, j s.t. $i \neq j$, any set of sequences $\mathbf{S}_i = (S_k)_{k=1}^M$ and $\mathbf{S}'_i = (S_k)_{k=1, k \neq i}^M \cap S'_i$ such that S_i and S'_i are neighboring and any subset of actions $S_j \subset \mathcal{D}_{j,1} \times \mathcal{D}_{j,2} \dots \times \mathcal{D}_{j,T}$ it holds that*

$$\mathbb{P}(A_j(\mathbf{S}_i) \in S_j) \leq \exp(\epsilon) \mathbb{P}(A_j(\mathbf{S}'_i) \in S_j) + \delta$$

Remark 2.1 (Item-level vs. User-level DP in Federated Learning). *Fed-JDP definition aims to protect each item from any of the users against the others. Thus, it is often called item-level DP and is widely studied in Federated bandits (Dubey and Pentland, 2020; Huang et al., 2021). Huang et al. (2023) proposes a stronger notion where privacy of the entire history is considered. This notion of user-level DP is out of the scope of present work.*

Remark 2.2 (Local DP vs. JDP in High-dimensions). *Wang et al. (2020); Zhou and Chowdhury (2024) propose to use local DP for low-dimensional linear contextual bandits due to its strong privacy-preserving properties. But in high-dimensional settings, it causes significant degradation of utility because non-private high-dimensional regression depends on $\Omega(\sqrt{s \log d})$ but high-dimensional regression with local DP depends on $\Omega(\sqrt{d})$. Here, s is the sparsity parameter and $s \ll d$. In contrast, high-dimensional regression with JDP shows $\Omega(\sqrt{s \log d})$ dependence (Chakraborty et al., 2024). Thus, we focus on JDP.*

High-dimensional Contextual Bandits. High-dimensional bandits with sparse linear structure are extensively studied (Bastani and Bayati, 2020; Oh et al., 2021; Hao et al., 2020). An algorithm for this problem follows a generic template. At every step t , (i) Observe context x_t , (ii) Construct an estimator, $\hat{\theta}_t$ from collected data, i.e. rewards and contexts $\{x_s, r_s\}_{s=1}^{t-1}$, (iii) Play arm greedily using $\hat{\theta}_t$ and observed context x_t , and (iv) Observe rewards r_t and store data. In this setting, the estimators of θ are often constructed using by solving the LASSO problem, i.e., $\min_{\theta, z} \frac{1}{t-1} \sum_{s=1}^{t-1} (r_s - \theta^\top x_s)^2 + \lambda \|z\|_1$ s.t. $\theta - zI_{d \times d} = 0$. For the private case, private thresholding-based mechanisms and Peeling algorithms Dwork et al. (2014b) are typically used. The analysis of these algorithms in centralized and federated, private and non-private settings differ significantly from each other motivating the following question: *Can we design an ADMM-based private LASSO algorithm, which is amenable to central, federated models with and without privacy for the high-dimensional bandit problem?*

Assumptions. We describe the assumptions that we use to establish regret guarantees of HiBPA. All of these assumptions are standard in high-dimensional bandit literature Zhang and Huang (2008); Zhang (2010); Bastani and Bayati (2020); Li et al. (2021); Chakraborty et al. (2023, 2024).

Assumption 2.1. (a) *Bounded context:* $\mathbb{P}_{x \sim P_i} (\|x\|_\infty \leq x_{\max}) = 1, \forall i \in [K]$ and $x_{\max} \in \mathbb{R}^+$.
(b) *Bounded parameters:* $\|\theta\|_0 \leq s_0$ and $\|\theta\|_1 \leq s_{\max}$.

Assumption 2.2 (Observational Noise). *We assume that the random variables ϵ_t are independent and each one is σ sub-Gaussian, i.e., $\mathbb{E}[\exp(\alpha \epsilon_t)] \leq \exp\left(\frac{\sigma^2 \alpha^2}{2}\right)$ for all $t \in [T]$ and $\alpha \in \mathbb{R}$.*

Definition 3 (s -sparse eigenvalues). *For a symmetric matrix A , its minimum and maximum s -sparse eigenvalues are defined as $\phi_{\min}(s, A) = \inf_{u: u \neq 0, \|u\|_0 \leq s} \frac{u^\top A u}{\|u\|_2^2}$, $\phi_{\max}(s, A) = \sup_{u: u \neq 0, \|u\|_0 \leq s} \frac{u^\top A u}{\|u\|_2^2}$.*

Assumption 2.3 (Context distribution). *The contexts corresponding to each arm i are generated stochastically from P_i that satisfies the following conditions.*

1. Bounded Orlicz norm: *For any arm $i \in [K]$, the Orlicz norm of context distribution P_i is bounded, i.e. $\|X_i\|_{\psi_2} \leq \nu$ for $X \sim P_i$.*
2. Sparsity: *For all $i \in [K]$, the design matrix $\Sigma_i = \mathbb{E}_{x \sim P_i} [xx^\top]$ has bounded maximum sparse eigenvalue, i.e. $\phi_{\max}(Cs^*, \Sigma_i) \leq \phi_u \leq \infty$.*

Assumption 2.4 (Margin condition). *There exists a positive constant Δ_* , $A, \alpha \in [0, \infty)$, such that for an $h \in [A\sqrt{\frac{\log d}{T}}, \Delta_*]$ and $\forall t \in [T]$, the following holds $\mathbb{P}\left(x_{k_t}^\top \theta \geq \max_{i \neq k} x_i^\top \theta + h\right) \leq \left(\frac{h}{\Delta_*}\right)^\alpha$.*

The zero-mean sub-Gaussian assumption on the noise (Assumption 2.2) is satisfied by various families of distributions, including normal distribution and bounded distributions, which are commonly chosen noise distributions. The margin condition (Assumption 2.4) controls the hardness of the bandit instance. For $\alpha \rightarrow \infty$, there is a deterministic gap between arms. $\alpha = 0$ implies that there is no apriori information about arm separation. The sparsity assumption (Assumption 3) on context distributions is needed to identify the sparse support of parameters (Chakraborty et al., 2023).

3 PeeledLASSO: Private ADMM with Peeling

In this section, we propose a template of designing PeeledLASSO algorithm for central and federated settings using ADMM-based optimizers (Boyd et al., 2011) and Peeling algorithm (Dwork et al., 2014b) in Algorithm 1. To obtain private LASSO estimates, we apply Peeling operator on each ADMM update. Introducing the dual variable $u = (u_1, \dots, u_N) \in \mathbb{R}^{N \times d}$ initialized to u_0 and exploiting the separable structure of the consensus problem, we obtain the updates for the i^{th} iteration of ADMM:

$$\hat{\theta}_i = \mathcal{P}_s \left(\frac{1}{N} \sum_{j=1}^N u_{i-1,j}, \sigma^2 \right), z_{i,j} = \text{prox}_{\gamma, \frac{1}{N}}(r_j - x_j^\top \cdot) \left(2\hat{\theta}_i - u_{i-1,j} \right), \quad (1)$$

$$u_{i,j} = u_{i-1,j} + 2\lambda(z_{i,j} - \hat{\theta}_{i+1}) \quad (2)$$

where, $\text{prox}_{\gamma, f}(v_0) = \arg \min_v \left\{ f(v) + \frac{1}{2\gamma} \|v - v_0\|_2^2 \right\}$. The updated variable θ is given by the peeling operator denoted by \mathcal{P}_s , which can be seen as a noisy projection on an ℓ_0 -ball of radii s ,

Algorithm 1 PeeledLASSO($\{(\mathbf{r}_{v,N}, \mathbf{X}_{v,N})\}_{v \in [M]}, \sigma^2, B, \lambda, \gamma, M, s$)

```

1: Input: Noise variance  $\sigma^2$ , #iterations  $B$ ,  $\gamma \in (0, 1]$ ,  $\lambda, s$ 
2: Initialize Collect  $\{u_{0,v}\}_{v=1}^M$  and set  $\hat{\theta}_1 = \frac{1}{M} \sum_{v=1}^M u_{0,v}$ 
3: for  $i = 1 : B$  do
4:   for  $v \in [M]$  do
5:      $u_{i,v} \leftarrow \text{LocalUpdate}(\mathbf{X}_v, \mathbf{r}_v, \theta_i, M, \lambda, \gamma)$  //User level computations//
6:   end for
7:    $\hat{\theta}_{i+1} \leftarrow \mathcal{P}_s(\frac{1}{M} \sum_{v \in [M]} u_{i+1,v}, \sigma^2)$  //Server level computations//
8:   Communicate  $\hat{\theta}_{i+1}$  to each user
9: end for
10: Return:  $\theta_B$ 

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Algorithm 2 LocalUpdate($\mathbf{X}_v = \{X_j\}_{j=1}^N, \mathbf{r}_v = \{r_j\}_{j=1}^N, \theta_i, M, \lambda, \gamma$)

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1: Parameters: Clipping parameter  $C = 4r_{\max} X_{\max}$ 
2: Initialize:  $u_{0,v}$ 
3:  $\Sigma_v \leftarrow (\frac{2\gamma}{MN} \mathbf{X}_v \mathbf{X}_v^\top + \mathbf{I}_d)$ 
4:  $z_{i,v} \leftarrow \Sigma_v^{-1} (2\gamma \mathbf{r}_v \mathbf{X}_v + u_{i,v} - 2\hat{\theta}_i)$ 
5:  $u_{i,v} \leftarrow u_{i-1,v} + 2\lambda (\text{Clip}(z_{i,v} - \hat{\theta}_i, C))$ 
6: Return:  $u_{i,v}$ 

```

where noise variance σ^2 is calibrated to ensure the desired privacy level. We refer to Algorithm 5 in Appendix B for details. From these updates and together with the possibility to randomly sample the blocks in our general scheme, we can naturally obtain different variants of ADMM for the centralized and federated models in their private and non-private counterparts. We observe this universality of ADMM scheme as lines 5-7 of Algorithm 1 extends to Algorithm 4. We further use this flexibility to construct the PeeledLASSO estimators for both federated and central settings.

Proposition 1 (Privacy Guarantees). *Under Assumptions 2.1 and 2.2, Algorithms 1 and 2 satisfy (ϵ, δ) -JDP (for $M = 1$) and (ϵ, δ) -FedJDP for $\sigma^2 = \frac{24s\sigma_\eta^2 \log MN \log(\frac{1}{\delta})}{M^2 N^2 \epsilon^2}$, where σ_η^2 is the variance of observational noise in the bandit model.*

Theorem 1 (Estimation Error). *We are given N samples and σ^2 is tuned as per Proposition 1 to ensure (ϵ, δ) -JDP and (ϵ, δ) -FedJDP in centralised and federated settings, respectively.*

(a) *In central setting, Algorithm 1 yields estimation error $\mathbb{E} [\|\hat{\theta}_B - \theta^*\|_2]$ of $\tilde{O} \left(\frac{s^* \log d}{\sqrt{N_\ell}} + \frac{s^* \log d}{N_\ell \epsilon} \right)$.*

(b) *In federated setting, Algorithm 1 yields estimation error $\mathbb{E} [\|\hat{\theta}_B - \theta^*\|_2]$ of $\tilde{O} \left(\frac{s^* \log d}{\sqrt{MN_\ell}} + \frac{s^* \log d}{\epsilon MN_\ell} \right)$.*

We emphasize the estimation error in Theorem 1 is established differently from that in antecedent work such as Bastani and Bayati (2020); Oh et al. (2021). In particular, we analyze ADMM as a fixed-point iteration viewing peeling and regression as proximal operators (see Appendix D for the complete proof).

Discussions. We see that in the central and non-private case, the estimation error decreases as $\frac{1}{\sqrt{N}}$ and $\frac{1}{\epsilon N}$ for the private case. Due to homogeneity of contexts, the estimation error scales as $\frac{1}{M}$ with the number of clients in the private case and scales $\frac{1}{\sqrt{M}}$ in the non-private case. It remains an interesting open problem to establish the tightness of these error bounds by establishing suitable lower bounds and is left as future work. Further, a critical observation here is that we are not adding any noise during the dual update (step 5 in Algorithm 4). While this works in the low-dimensional setting, it would not lead to good utility guarantees for the high-dimensional problem since the noise variance would be proportional to d leading to the utility decay as $O(T)$. Further, we would like to remark that while the bounds presented in Theorem 1 are under the ℓ_2 norm, since the sparsity of θ is

Algorithm 3 HiBPA: High-dimensional Bandits with Peeled ADMM

```

1: Input: Privacy level  $\epsilon, \delta$ , users  $[M]$ , sparsity upper bound  $s, \gamma$ 
2: Require: bounds  $b_{\max}, X_{\max}$ , sub-gaussian parameter  $\sigma_\eta^2$ 
3: Initialize:  $t_0 = 0, \ell = 0, \lambda_0$ 
4: if  $\ell = 0$  then
5:   for user  $v \in [M]$  do
6:     Play random arm  $k_{v,1} \in [K]$ , and observe  $r_{v,\ell}$ 
7:     Set  $\mathbf{X}_{v,0} = \{X_{v,k_{v,1}}\}, \mathbf{r}_{v,0} = \{r_{v,1}\}, t_1 = 1$ 
8:   end for
9: end if
10: for  $\ell = 1, 2, \dots$  do
11:   Set  $\sigma_\ell^2 = \frac{24s\sigma_\eta^2 \log(M(t_\ell - t_{\ell-1})) \log(\frac{1}{\delta})}{M^2(t_\ell - t_{\ell-1})^2 \epsilon^2}$ , clipping levels  $C_\ell$ , #iterations  $B_\ell$ , reg.  $\gamma_\ell$ , dual step  $\lambda_\ell$ 
12:    $\hat{\theta}_\ell \leftarrow \text{PeeledLASSO}(\{\mathbf{X}_{v,\ell-1}, \mathbf{r}_{v,\ell-1}\}_v, \sigma_\ell^2, B_\ell, C_\ell, \lambda_\ell, \gamma, M)$ 
13:   Set  $t_{\ell+1} = 2^{\ell+1}$ 
14:   for user  $v \in [M]$  do
15:      $(\mathbf{X}_{v,\ell}, \mathbf{r}_{v,\ell}) \leftarrow \text{CollectBanditData}(t_{\ell+1} - t_\ell, \hat{\theta}_\ell, v)$ 
16:   end for
17: end for

```

201 bounded by s^* , these bounds can be translated in terms of s (i.e., independent of d). This fact will be
202 used in the proof of Theorem 3, 4.

203 **Remark 3.1** (Choices of Adding Noise). *The noise to ensure privacy could be added at different*
204 *points of communication, e.g. Step 5, 8, or 9 in Algorithm 1. We chose to add it in Step 5 after*
205 *updating θ_i at each iteration i . This is a conscious choice as the ADMM iteration can be stopped*
206 *at any time by a practitioner and in federated setting, θ_i is communicated to each user after each*
207 *update. If we add noise in the other two steps, we have to calibrate it across each user and each*
208 *data point. This leads to sampling noise more number of times. Adding it to θ_i minimizes the process*
209 *while ensuring both privacy and sparsity.*

210 **Remark 3.2** (Peeling vs. Soft Thresholding). *In some variants of private LASSO with ADMM,*
211 *soft thresholding is used in Step 5 of Algorithm 1 and Step 11 of Algorithm 1. Though this works*
212 *perfectly in offline setting (Cyffers et al., 2023), in our bandit setting, we need to control the sparsity*
213 *of PeeledLASSO estimate at any point of time (Chakraborty et al., 2024; Shukla, 2024). This is*
214 *essential to ensure eventual recovery of support of the true parameter, and thus, correctness of the*
215 *final bandit algorithm. While soft-thresholding does not exhibit this property, we use Peeling to*
216 *ensure these requirements.*

217 4 Algorithm Deisgm: High-dimensional Bandits with PeeledLASSO

218 In this section, we present our main algorithm for privately learning high-dimensional bandits,
219 (Algorithm 3). HiBPA is an episodic algorithm that computes θ_t^{priv} privately using ADMM updates.
220 We divide the decision horizon into episodes of geometrically progressive lengths. Specifically, the
221 ℓ^{th} episode begins at time 2^ℓ and the underlying algorithm restarts. In the private case, the doubling
222 trick helps us reduce the amount of noise needed to preserve the privacy of the data. Since there are
223 only logarithmic number of episodes. At each time t , the algorithm observes the set of context vectors
224 \mathcal{X}_t using $\hat{\theta}_t^{\text{priv}}$. The estimator $\hat{\theta}_t^{\text{priv}}$ is computed depending on the communication protocol used. Using
225 this estimator, it selects an arm greedily at the current time-step. For both these subroutines, we use
226 the forgetfulness technique, wherein the data collected *only* from the previous episodes is used to
227 compute the estimates of the current episode. Given the estimator, we play an arm greedily using
228 $\text{CollectBanditData}(N, \hat{\theta}, v)$ and collect resulting data.

229 4.1 Central model

230 In the centralized model, a trusted curator holds the dataset $\{(X_s, r_s)\}_{s=1}^t$ and creates a private
231 estimator from this dataset. Our private ADMM algorithm for this centralized model follows the
232 updates in (1) the CentralOPLASSO subroutine. The CentralOPLASSO sub-routine takes as input a
233 random vector and uses initial data to compute admm-updates. These u -variable updates are then

Algorithm 4 CollectBanditData($N, \hat{\theta}, v$)

```
1: Initialize  $\mathbf{X}_{v,\ell} = \emptyset, \mathbf{r}_{v,\ell} = \emptyset$ 
2: for steps  $t = 1, \dots, N$  do
3:   Play  $k_{v,t} = \arg \max_{k \in [K]} x_{v,k}^\top \hat{\theta}$ 
4:   Observe reward:  $r_{v,t} = X_{v,k_{v,t}}^\top \theta + \eta_t$ 
5:    $\mathbf{X}_{v,\ell} = \mathbf{X}_{v,\ell} \cup \{X_{v,k_{v,t}}\}, \mathbf{r}_{v,\ell} = \mathbf{r}_{v,\ell} \cup \{r_{v,t}\}$ 
6: end for
7: Return
```

234 peeled *every iteration* to compute the final θ update. One could add it after completing a pre-decided
235 number of iterations but in practice one can stop the optimization prematurely. In this case, if we
236 publish and use the estimated parameter, it will lead to privacy leakage. Thus, we claim to add it after
237 each iteration. Additionally, as we will see in the next section, this is a mandatory requirement in
238 federated model. Thus, peeling the estimated parameter after each iteration allows us to formulated a
239 unified algorithmic template and analysis without losing anything in utility.

240 4.2 Federated model

241 For the Federated Learning (FL) model Kairouz et al. (2021), we assume that there exists a central
242 server that coordinates the behaviors of all the different clients. The server has access to the same
243 partition of the parameter space used by all the clients, and can communicate with the clients. Every
244 client is a K -armed high-dimensional contextual bandit problem, where the true parameter is shared
245 across clients. Due to privacy concerns, the client-side algorithm should keep the reward of each
246 evaluation confidential. The only things that can be transmitted to the server are the local statistical
247 summary of the rewards. The clients are not allowed to communicate with each other. Accordingly,
248 the FedOPLASSO subroutine samples a set of users and uses local data to perform admm-updtes. Then,
249 the u -variables from every user are aggregated and peeled in order to obtain the global theta updates.
250 These updates are executed in a federated fashion since (i) the blocks x_k and u_k associated to each
251 arm k can be updated and perturbed locally and in parallel, and (ii) if each arm k shares $u_{t+1} - u_t$
252 with the server, then the latter can execute the rest of the updates to compute z_{t+1} . On top of this
253 vanilla version, we can natively accommodate user sampling (often called “client sampling” in the
254 literature), which is a key property for cross-device FL as it allows to improve efficiency and to model
255 partial user availability Kairouz et al. (2021). The communication cost of such updates is $Md \log T$
256 over T episodes if all users participate at every step. Finally, in the federated setup it is important to
257 note that parameters λ, γ are tuned centrally not locally since the final updates happen centrally.

258 **Numerical Experiments:** We compare HiBA with existing baselines and show that it performs
259 competitively in settings considered in the paper. Due to space constraints all our experimental studies
260 are deferred to Appendix ??.

261 5 Privacy and Utility Guarantees of HiBPA

262 Now, we establish privacy and utility guarantees for HiBPA in both central and federated models.

263 **Theorem 2** (Privacy Guarantees). *Under Assumption 2.1 and 2.2, Algorithm 3 preserves (ϵ, δ) -JDP*
264 *and (ϵ, δ) -Fed-JDP in the central and federated models, respectively.*

265 HiBPA satisfies due to two components. First, the parameters estimated in each episode ℓ using
266 CentralOPLASSO and FedOPLASSO are already (ϵ, δ) -JDP and (ϵ, δ) -Fed-JDP due to Proposition 1.
267 Second component is forgetting with doubling episodes as proposed by (Azize and Basu, 2022).
268 Since estimated parameters are the only thing communicated to the server in each episode and are
269 used for further data collection, ensuring their privacy protects the data collected in the previous
270 episode. Additionally, since we do not reuse data of any past episode except the last one, their privacy
271 is not leaked any further. Additionally, HiBPA also enjoys reward-DP ensuring privacy protection for
272 collected downstream rewards. In the reward-DP model, the adversary is only allowed to perturb the
273 reward stream (Hanna et al., 2024; Azize and Basu, 2024). In Theorem 3 and 4, we establish utility
274 guarantees for HiBPA under both central and federated models, and also with and without (ϵ, δ) -JDP.

275 **Theorem 3** (Central model). *Let us set $C_\ell = x_{\max} b_{\max} + 2\sigma \sqrt{\log N_\ell}$, $\lambda_\ell = \frac{1}{1-\sqrt{\gamma_\ell}}$.*

276 1. *Non-private:* Under Assumption 2.1, 2.2, 2.3, and 2.4, Algorithm 3 with parameters $\gamma_\ell =$
 277 $\sqrt{\frac{\log d N_\ell}{N_\ell}}, B_\ell = \mathcal{O}\left(\log_{1/(1+\gamma_\ell)}(N_\ell)\right)$ incurs a regret

$$\mathbb{E}[R(T)] \leq x_{\max} b_{\max} s \log d + \psi(\alpha),$$

278 such that $\psi(\alpha) = \frac{T^{(1-\alpha)/2}}{(1-\alpha)\Delta_*^\alpha} (s^3 \log d)^{\frac{1+\alpha}{2}}, \alpha \in [0, \infty), s^{3/2} \log d, \alpha = \infty.$

279 2. *Private:* Under assumptions 2.1, 2.2, 2.3, and 2.4, Algorithm 3 with parameters $\gamma_\ell =$
 280 $\sqrt{\frac{\log d N_\ell}{N_\ell}}, B_\ell = \mathcal{O}\left(\log_{1/(1+\gamma_\ell)}(N_\ell \epsilon^2)\right), \sigma_\ell^2 = \frac{24s\sigma_\eta^2 \log(\frac{N_\ell}{\delta})}{(N_\ell \epsilon)^2}$ incurs a regret

$$\mathbb{E}[R(T)] \leq x_{\max} b_{\max} s \log d + \psi(\alpha) \wedge \phi(\epsilon),$$

281 such that $\psi(\alpha)$ is as above and $\phi(\epsilon) = \Psi_\alpha(T) \left(\frac{s^3 \log^2 d \log(\frac{1}{\delta})}{\epsilon^2} \right)^{\frac{(1+\alpha)}{2}}, \alpha \in [0, \infty),$ and
 282 $\frac{s^3 \log^2 d \log(\frac{1}{\delta})}{\epsilon^2}, \alpha = \infty.$ For $\alpha > 0, \Psi_\alpha(T) \lesssim \frac{\Gamma(3+2\alpha)}{\Delta_*^\alpha (\alpha \log 2)^{3+2\alpha}},$ and $\Psi_0(T) \lesssim \log^3 T.$

283 **Discussions: Implications of Theorem 3.** 1. *Dependence on dimension and sparsity:* We observe
 284 that with respect to the dimension the privacy term dominates the non-privacy term by a factor
 285 of $(\log d)^{1+\alpha/2}$. In contrast, they depend similarly on the sparsity parameter $s^* + s$, i.e. $(s +$
 286 $s^*)^{\frac{3(1+\alpha)}{2}}$. Finally, note that the regularization parameter is independent of the privacy level since in
 287 PeeledLASSO privacy is achieved through the peeling step.

288 2. *Dependence on T :* The non-private depends on $T^{-\frac{1+\alpha}{2}}$ whereas the private part has approximately
 289 $\log^3 T$ dependence. Thus, as $T \rightarrow \infty$, the effect of T fades away. Additionally, for $\alpha > 0$, the private
 290 term of regret is T independent. The last two observations resonate with the similar impact of privacy
 291 in other bandit problem such as finite-armed (Azize and Basu, 2022), linear (Azize and Basu, 2024),
 292 and stochastic low-dimensional contextual bandits (Azize and Basu, 2024; Hanna et al., 2022).

293 3. *Minimax optimality:* For $\alpha = 0$, Chakraborty et al. (2024) provide the lower bound for (ϵ, δ) -JDP
 294 high-dimensional linear bandits in the central model. We observe that our regret is *order-optimal in*
 295 *terms of T and dimension both with and without privacy.* For sparsity, the non-private and private
 296 regret bounds has an additional multiplicative $s + s^*$ factor. Now, we state the regret upper bounds of
 297 HiBPA for non-private and private federated model.

298 4. *Regularization for Exploration.* Here, we adaptively tune the regularization parameter of LASSO
 299 as $\sqrt{\frac{\log d N_\ell}{N_\ell}}$. This is commonly used in high-dimensional bandits for enforcing exploration (Oh et al.,
 300 2021; Ariu et al., 2020).

301 **Theorem 4** (Federated model).

302 1. *Non-private:* Under assumptions 2.1, 2.2, 2.3, and 2.4, Algorithm 3 with parameters $\gamma_\ell =$
 303 $\sqrt{\frac{\log d N_\ell}{N_\ell}}, B_\ell = \mathcal{O}\left(\log_{1/(1+\gamma_\ell)}(MN_\ell)\right)$ incurs a regret

$$\mathbb{E}[R(T)] \leq x_{\max} b_{\max} s \log d + \psi(\alpha),$$

304 such that $\psi(\alpha) = \frac{(MT)^{(1-\alpha)/2}}{(1-\alpha)\Delta_*^\alpha} (s^3 \log d)^{\frac{1+\alpha}{2}}, \alpha \in [0, \infty),$ and $s^3 \log d, \alpha = \infty.$

305 2. *Private:* Under assumptions 2.1, 2.2, 2.3, and 2.4, Algorithm 3 with parameters $B_\ell =$
 306 $\mathcal{O}\left(\log_{1/(1+\gamma_\ell)}(MN_\ell \epsilon^2)\right), \sigma_\ell^2 = \frac{24s\sigma_\eta^2 \log(MN_\ell) \log(\frac{1}{\delta})}{(MN_\ell \epsilon)^2}$ incurs a regret

$$\mathbb{E}[R(T)] \leq x_{\max} b_{\max} s \log d + \psi(\alpha) \wedge \phi(\epsilon),$$

307 such that $\psi(\alpha)$ is as above and $\phi(\epsilon) = \Psi_\alpha(MT) \left(\frac{s^3 \log^2 d \log(\frac{1}{\delta})}{\epsilon^2} \right)^{\frac{(1+\alpha)}{2}}, \alpha \in [0, \infty),$ and
 308 $\frac{s^3 \log^2 d \log(\frac{1}{\delta})}{\epsilon^2}, \alpha = \infty.$

309 **Discussion: Dependence of HiBPA on M .** We observe that without privacy regret of HiBPA in
 310 federated model varies with $M^{\frac{1-\alpha}{2}}$. For $\alpha = 0$, it is \sqrt{M} and fades away as $M \rightarrow \infty$. In contrast,

due to the nature of Ψ_α being a bounded function, the privacy term of regret is independent of M for $\alpha > 0$. For $\alpha = 0$, it depends on $\log^3 M$. In the next section, we prove a lower bound on regret of federated bandit algorithms for stochastic high-dimensional linear contextual bandits satisfying (ϵ, δ) -Fed-JDP. The lower bounds indicate that non-private regret of HiBPA achieves order-optimal dependence on M for $\alpha = 0$, whereas it is loose by a $\log^3 M$ factor for regret term due to privacy.

6 Lower Bound

In this section, we report the cost of privacy in the high-dimensional bandit in the federated model with privacy constraints. We propose a lower bound on the regret of any (ϵ, δ) -JDP algorithm for the high-dimensional bandit problem. This will shed further light on the cost of privacy in these settings when compared with other existing non-private lower bounds.

Definition 4 (Minimax regret). *We define the minimax regret in the (ϵ, δ) -JDP setting as:*

$$R_{(\epsilon, \delta)}^{\text{minimax}} := \inf_{\pi \in \Pi_{(\epsilon, \delta)}} \sup_{P_{x, r} \in \mathcal{B}} \mathbb{E}[R(T)],$$

where \mathcal{B} denotes the set of all sparse high-dimensional bandit instances satisfying Assumptions 2.1 and 2.2, and $\Pi_{(\epsilon, \delta)}$ denotes the set of all collinearly dependent (ϵ, δ) -Fed-JDP policies.

Following Huang et al. (2021), we consider the collinear policies to rigorously proof the lower bound. Intuitively, if two clients that are not collinear under a policy, their local observations on any arm cannot be utilized to improve the others' knowledge about the arm. Thus, it is equivalent to running M independent bandits. The policy used by HiBPA to interact is also a collinear policy.

Theorem 5 (Lower bound: Federated, Private). *Give homogeneous context distributions across users, privacy parameters $\epsilon, \delta > 0$ such that $\epsilon^2 < \log(\frac{1}{\delta})$, and $d > 9$ and sufficiently large $s \log(d/s)$, we have that*

$$R_{(\epsilon, \delta)}^{\text{minimax}}(T) = \Omega \left(\max \left\{ \underbrace{\sqrt{\frac{s^* \log^2(d/s^*) \log(1/\delta)}{\epsilon^2}}}_{\text{private and high-d}}, \underbrace{\sqrt{s^* MT \log(d/s^*)}}_{\text{non-private and high-d}} \right\} \right).$$

Discussions: 1. *Impact of M .* This shows that for federated bandit with M users with homogeneous context distributions the non-private regret scales with $\tilde{\Omega}(\sqrt{M})$. HiBPA achieves this scaling w.r.t. M . In low-dimensional federated bandits settings, \sqrt{M} dependence is exhibited by (Zhou and Chowdhury, 2024) but there exists no lower bound to support this phenomenon. Our lower bound shows that \sqrt{M} is the optimal dependence.

2. *Private vs. Non-private Bound.* In contrast to the non-private lower bound, the private lower bound does not depend on M and T . This indicates a regime change depending on the privacy budget ϵ and the non-private bound dominates for $\epsilon \geq \sqrt{\frac{\log d}{MT}}$. For $M = 1$, we retrieve the lower bound of (Chakraborty et al., 2024) for sparse high-dimensional bandits in centralised setting.

3. *Context Homogeneity.* Homogeneity of contexts is fundamentally important in this lower bound. Since we collect sample from M users at each step and their context distributions are homogeneous, conceptually we can treat them together as MT samples from the bandit environment. This intuition is reflected in the lower bound. Though homogeneity is prevalent in federated bandit literature (Dubey and Pentland, 2020; Wang et al., 2023, 2020), context heterogeneity is gaining interest due to its practicality (Blaser et al., 2024). It is still an open problem to develop lower bounds for federated bandits with heterogeneous users and we leave it as a future work.

7 Conclusion

In this paper, we considered the high-dimensional bandit problem under different communication models and privacy constraints. We proposed a unifying algorithmic design and analysis template that can be used for all these models. This is accompanied by a novel privacy and utility analysis of this framework. We also show that our performance matches the lower bounds up to a factor of s . As a future work, it would be interesting to extend this problem to the case where the clients are heterogeneous.

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Part I

Appendix

A Related Work

A.1 Existing Literature

Multi-armed bandits have been studied since the foundational work of (Lai et al., 1985) and a comprehensive overview can be found in (Lattimore and Szepesvári, 2020). In this work, we consider the high-dimensional bandit problem (Li et al., 2021; Hao et al., 2020; Carpentier and Munos, 2012), under the necessary assumption that the unknown parameter θ is s^* sparse. (Bastani and Bayati, 2020; Oh et al., 2021; Ariu et al., 2020) propose LASSO or thresholding-based algorithms for this problem with competitive regret guarantees. Extending this line work to ensure privacy requires constructing estimators of the mean reward by injecting well-calibrated noise. This however implies that the variance of noise – and therefore the error in the constructed estimate – scales linearly with the dimension of the parameter. A straightforward adaptation of these tools to the high-dimensional problem setting where $d \gg T$ would lead to a regret bound that scales $\mathcal{O}(d)$, which would implicitly be super-linear in T . Another approach to designing private algorithms is objective or gradient perturbation (Chaudhuri et al., 2011; Kifer et al., 2012). (Talwar et al., 2015; Asi et al., 2021; Cyffers et al., 2023; Raff et al., 2024) consider this problem in the offline setting.

Federated learning (FL) (McMahan et al., 2017) has become a popular distributed machine learning paradigm in which numerous clients collaboratively train a prediction model under the coordination of a central server while maintaining the local training data at each client. FL is motivated by various applications where real-world data are exogenously generated at edge devices, and it is desirable to protect the privacy of local data by only sharing model updates instead of the raw data. In the low-dimensional setting, one of the first works to consider Federated linear contextual bandits was (Huang et al., 2021) and (Dubey and Pentland, 2020) extended this to a Fed-DP guarantee for their algorithm.

B Peeling

Algorithm 5 Peeling

```

1: Input: Vector  $v \in \mathbb{R}^d$ , sparsity  $s$ , privacy level  $\epsilon, \delta$ , sensitivity  $D$ 
2: Initialise:  $S = \emptyset$  and set  $\xi = \frac{2D\sqrt{3s \log(1/\delta)}}{\epsilon}$ 
3: for  $i = 1, 2, \dots, s$  do
4:   Generate  $\mathbf{w}_i = (w_{i1}, w_{i2}, \dots, w_{id}) \stackrel{iid}{\sim} \text{Lap}(\xi)$ 
5:    $j_i^* = \arg \max_{j \in [d] \setminus S} |v_j| + w_{ij}$ 
6:   Update  $S \leftarrow S \cup \{j_i^*\}$ 
7: end for
8: Set  $\mathcal{P}_s = v_S$ 
9: Generate  $\tilde{\mathbf{w}} = (w_{i1}, w_{i2}, \dots, w_{id}) \stackrel{iid}{\sim} \text{Lap}(\xi)$ 
10: Return:  $\mathcal{P}_s(v; \epsilon, \delta, D) = \mathcal{P}_s(v) + \tilde{\mathbf{w}}_S$ 

```

Lemma 1 ((ϵ, δ)-DP Dwork et al. (2018)). *If for every pair of adjacent datasets \bar{D}, \bar{D}' , we have $\|v(\bar{D}) - v(\bar{D}')\| \leq D$, then Algorithm 5 is (ϵ, δ) -DP.*

We further have the following bounds on that establish that peeling is indeed a contraction Cai et al. (2020).

Lemma 2 (Peeling). *For any index set $I \subset [d]$ and $\theta \in \mathbb{R}^d$ with $\text{supp}(\theta) \subset I$ and any $\hat{\theta} \in \mathbb{R}^d$ such that $\|\theta\|_0 \leq s$, we have for every $c > 0$, we have:*

$$\|\hat{\theta} - \theta\|_2^2 \leq \left(1 + \frac{1}{c}\right) \frac{I - s}{I - \hat{s}} \|\hat{\theta} - \theta\|_2^2 + 4(1 + c) \sum_i \|w_i\|_\infty$$

820 **Lemma 3** (Peeling is a contraction). *Given $\theta^* \in \mathbb{R}^d$ with $\|\theta^*\|_0 \leq s$ and θ such that $\text{supp}(\theta^*) \subset$
821 $\text{supp}(\theta)$ with $\|\theta\|_0 \leq \tilde{s} + s$, for any $c_1, c_2 > 0$ and $c > 0$, we have:*

$$\|\mathcal{P}_s(\theta, W, \tilde{w}) - \theta^*\|_2^2 \leq a \|\theta - \theta^*\|_2^2 + b$$

822 *where,*

$$\begin{aligned} a &= \left(1 + \frac{1}{c_1}\right) \left[\left(1 + \frac{1}{c_2}\right) \left(1 + \frac{1}{c}\right) + (1 + c_2) \right] \\ b &= (1 + c_1) \|\tilde{w}_S\|_2^2 + 4 \left(1 + \frac{1}{c_1}\right) \left(1 + \frac{1}{c_2}\right) (1 + c) \sum_{i \in [s]} \|w_i\|_\infty^2 \end{aligned}$$

C Preliminaries

Many optimization algorithms such as proximal methods, ADMM belong to this family Bauschke and Combettes (2019). Formulating optimization algorithms as λ -averaged operators allows us to use generic convergence results by the Krasnosel'skii Mann theorem Byrne (2003). In the setting of differential privacy, each application of T is perturbed by noise due to stochasticity in the data. A generic convergence analysis of fixed-point iterations under both inexact and block updates Combettes and Pesquet (2021).

C.1 ADMM

Consider the problem of minimizing a function $f : \mathcal{X} \rightarrow \mathbb{R}$ where $\mathcal{X} \subseteq \mathbb{R}^p$. The problem reduces to finding a fixed-point, $x^* \in \mathcal{X}$, such that $0 \in \partial f(x^*)$ or $\nabla f(x^*) = 0$ under differentiability assumptions. Alternatively, we can view these updates as reaching the fixed point of an operator $T : \mathcal{X} \rightarrow \mathcal{X}$. Starting with initial point x_0 , these updates are given by:

$$x_{t+1} = T(x_t)$$

We now define the type of operators considered in this paper.

Definition 5 (Fixed Point Operator). *Let $T : \mathcal{X} \rightarrow \mathcal{X}$ and $\lambda \in [0, 1]$, then:*

1. *T is non-expansive, i.e., $\|T(x) - T(x')\| \leq \|x - x'\|$, $\forall x, x' \in \mathcal{X}$. It is τ contractive if it is τ -Lipschitz.*
2. *T is λ -averaged if there exists a contractive operator R such that $T = \lambda R + (1 - \lambda)I$.*

We now present how ADMM can be defined as a fixed-point iteration. ADMM minimizes the sum of two (possibly non-smooth) convex functions with linear constraints between the variables of these functions, which can be formulated as:

$$\begin{aligned} \min_{x, z} \quad & f(x) + g(z) \\ \text{s.t.} \quad & Ax + Bz = c \end{aligned}$$

ADMM is often presented as an approximate version of the augmented Lagrangian method, where the minimization of the sum in the primal is approximated by the alternating minimizations on x and z . However, for the purpose of analysis, it is useful to view ADMM as a splitting algorithm Eckstein and Yao (2015). i.e., an approach to find a fixed point of the composition of two (proximal) operators by performing operations that involve each operator separately.

ADMM is defined through the Lions-Mercier operator Lions and Mercier (1979). Given two proximable functions P_1 and P_2 and parameter $\eta > 0$, the Lions-Mercier operator is: $T_{\eta P_1, \eta P_2} = \lambda R_{P_1} \eta R_{P_2} + (1 - \lambda)I$, where, $R_{P_1} = 2\text{prox}(P_1) - I$, $R_{P_2} = 2\text{prox}(P_2) - I$. In this work, we study fixed-point iterations with Differential Privacy (DP) Dwork et al. (2014a). DP relies on a notion of neighboring datasets. We denote a private dataset of size n by $D := (d_1, \dots, d_n)$. Two datasets D, D' are neighboring if they differ in at most one element. We refer to each d_i as a data item. Consider the problems of the form:

$$\min_{x \in \mathcal{X}} \frac{1}{n} \sum_{i=1}^n f(u; d_i) + r(u), \quad i = 1, \dots, n \quad (3)$$

where $f(\cdot; d_i)$ is a (typically smooth) loss function computed on data item d_i and r is a (typically non-smooth) regularizer. We denote $f(u; D) := \frac{1}{n} \sum_{i=1}^n f(u; d_i)$, $i = 1, \dots, n$. To solve this problem, we consider the general noisy fixed-point iteration described in Algorithm 1. The core of each update applies a λ_k -averaged operator constructed from a non-expansive operator R , and a Gaussian noise term added to ensure differential privacy via the Gaussian mechanism. Algorithm 1 can use (possibly randomized) block-wise updates ($B > 1$) and accommodate additional errors in operator evaluation.

861 D Proof of Theorem 1

862 *Proof.* For every iteration b given by (1) from Lemma 3, we have:

$$\begin{aligned} \mathbb{E} \left[\|\hat{\theta}_b - \theta^*\|_2^2 \right] &= \mathbb{E} \left[\|\mathcal{P}_S(\hat{\theta}_{b+0.5}, \mathbf{W}^{(b)}, \tilde{w}^b) - \theta^*\|_2^2 \right] \\ &\leq \mathbb{E} \left[\|\mathcal{P}_S((\hat{\theta}_{b+0.5})_{S_b}, \mathbf{W}^{(b)}, \tilde{w}^b) - \theta^*\|_2^2 \right] \\ &\leq c_3 \mathbb{E} \left[\|(\hat{\theta}_{b+0.5})_{S_b} - \theta^*\|_2^2 \right] + c_4 \\ &\leq c_3 \mathbb{E} [\|u_b - u^*\|_2] + c_4, \end{aligned}$$

863 where,

$$\begin{aligned} c_3 &= \frac{6}{7} \left(1 + \frac{1}{c_1} \right) \left[\left(1 + \frac{1}{c_2} \right) \left(1 + \frac{1}{c} \right) \frac{s^*}{s} + (1 + c_2) \right] \\ c_4 &= (1 + c_1) \mathbb{E} [\|\tilde{\mathbf{w}}_S\|_2^2] + 4 \left(1 + \frac{1}{c_1} \right) \left(1 + \frac{1}{c_2} \right) (1 + c) \mathbb{E} \left[\sum_{i \in [s]} \|\mathbf{w}_i\|_\infty^2 \right] \end{aligned}$$

864 Now we analyse the convergence step for which we use the analysis from ADMM. For the soft-
865 thresholding operator we have $\tau = \frac{1}{\gamma}$ and $D =$ and B_ℓ is the number of iterations for which ADMM
866 is run in a single episode. From Algorithm 4 and using Theorem 6, we have that:

$$\mathbb{E} [\|u_b - u^*\|_2^2] \leq \left(1 - \frac{(1 - \tau)}{8} \right)^{B_\ell} D$$

867 Plugging this back, we get:

$$\mathbb{E} [\|\hat{\theta}_{B_\ell} - \theta^*\|_2^2] \leq c_3 \left(1 - \frac{(1 - \tau)}{8} \right)^{B_\ell} D + c_4 B_\ell \mathbb{E} \left[\|\tilde{w}_{S_m}^{(m)}\|_2^2 + \sum_{i \in [s]} \|\mathbf{w}_i^{(m)}\|_\infty^2 \right]$$

868 The first step in Algorithm 1 is Peeling Algorithm 5. The noise added at every iteratino in order to
869 preserve the privacy is given by $\mathbf{W}_b = \left(\sum_{i \in [s]} \|\mathbf{w}_i^{(b)}\|_\infty^2 + \|\tilde{\mathbf{w}}_S^{(b)}\|_2^2 \right)$ where each $\mathbf{w}_i^{(b)} \in \mathbb{R}^d$ and

870 $\tilde{\mathbf{w}}_S$ has $|S| = s$ non-zero co-ordinates. Noting that $\xi = \frac{\theta_{\max}}{\kappa} \frac{\sqrt{3s \log \frac{B_\ell}{\delta}}}{N_\ell \epsilon}$, using Theorem 7, for
871 suitably large constant, we have:

$$\mathbb{P} \left(\|\mathbf{w}_S\|_\infty^2 \geq \frac{s \log^2 d \left(\log \frac{B_\ell}{\delta} \right)}{N_\ell^2 \epsilon^2} \right) \leq \frac{1}{d^8}$$

872 Using a union bound, we get:

$$\begin{aligned} &\mathbb{P} \left(\max_{b \in [B_\ell]} \mathbf{W}_b > K \frac{s^2 \log^2 d \left(\log \frac{B_\ell}{\delta} \right)}{N_\ell^2 \epsilon^2} \right) \\ &\leq \sum_{b \in [B_\ell]} \mathbb{P} \left(\mathbf{W}_b > \frac{s^2 \log^2 d \left(\log \frac{B_\ell}{\delta} \right)}{N_\ell^2 \epsilon^2} \right) \\ &\leq \frac{2sB_\ell}{d^8} = O \left(\frac{1}{d^6} \right) \end{aligned}$$

873 Plugging this back, we get:

$$\begin{aligned} \mathbb{E} [\|\hat{\theta}_B - \theta^*\|_2^2] &\leq c_3 \left(1 - \frac{(1 - \gamma)}{8} \right)^{B_\ell} D + c_4 B_\ell \mathbb{E} \left(\|\tilde{w}_{S_m}^{(m)}\|_2^2 + \sum_{i \in [s]} \|\mathbf{w}_i^{(m)}\|_\infty^2 \right) \\ &\leq c_3 \left(1 - \frac{(1 - \gamma)}{8} \right)^{B_\ell} D + c_4 B_\ell K' \cdot \frac{(s^* \log d)^2 \log \frac{\log B_\ell}{\delta}}{N_\ell^2 \epsilon^2} \end{aligned}$$

874

1. For the central setting, let B such that

$$\left(1 - \frac{1-\gamma}{8}\right)^{B_\ell} D = K' B_\ell \cdot \frac{(s^*)^2 \log^2 d \log \frac{B_\ell}{\delta}}{N_\ell \epsilon^2}$$

$$B_\ell = \mathcal{O}\left(\log_{1/(1+\gamma)}\left(\frac{N_\ell \epsilon}{(s^* \log d)^2}\right)\right).$$

875

Ignoring log and log log terms, this leads to:

$$\mathbb{E}\left[\|\hat{\theta}_B - \theta^*\|_2^2\right] \lesssim C^2 \cdot \frac{(s^* \log d)^2}{N_\ell} \left(1 + \frac{1}{\epsilon^2 N_\ell}\right)$$

$$\implies \mathbb{E}\left[\|\hat{\theta}_B - \theta^*\|_2\right] \lesssim C \left(\frac{s^* \log d}{\sqrt{N_\ell}} + \frac{s^* \log d}{N_\ell \epsilon}\right).$$

876

2. In the Federated setting,

$$\mathbb{E}\left[\|\hat{\theta}_B - \theta^*\|_2^2\right] \leq c_3 \left(1 - \frac{(1-\gamma)}{8}\right)^{B_\ell} D + c_4 B_\ell K' \cdot \frac{(s^* \log d)^2 \log \frac{\log B_\ell}{\delta}}{M^2 N_\ell^2 \epsilon^2}$$

$$B = \mathcal{O}\left(\log_{1/(1+\gamma)}(MN_\ell)\right).$$

877

Leading to the following bound error per-iteration:

$$\mathbb{E}\left[\|\hat{\theta}_B - \theta^*\|_2\right] \lesssim \frac{s^* \log d}{\sqrt{MN_\ell}} + K' \cdot \frac{s^* \log d}{\epsilon MN_\ell}.$$

878

□

E Proof of Theorem 2

Proof. Federated Model: Let $\mathbf{r}_t = (r_{1,t}, r_{2,t}, \dots, r_{MK,t})$ and $\mathcal{X}_t = (X_{1,t}, \dots, X_{MK,t})$. Consider two neighboring datasets $\mathcal{D}_1 = \{(\mathbf{r}_t, \mathcal{X}_t)\}_{t \in [T]}$, $\mathcal{D}_2 = \{(\mathbf{r}'_t, \mathcal{X}'_t)\}_{t \in [T]}$. Let the randomized mapping introduced by Algorithm 1 be \mathcal{M}_t . At time t , \mathcal{M}_t takes a set of observed reward-context pairs from the previous episodes and recommends action $k_t \in [K]$. Let $\alpha(t)$ denote the episode to which time t belongs:

$$\alpha(t) := \sum_{\ell=0}^{\lfloor \log_2 T \rfloor} \ell \mathbf{1}\{t_\ell \leq t \leq t_{\ell+1}\}$$

$$\mathbb{P}(\mathcal{M}_\tau = k_{-\tau}) = \prod_{\ell \geq 0} \mathbb{P}(k_t = k_t; t_\ell \leq t \leq t_{\ell+1}, t \neq \tau)$$

Now, we will analyze this product. First note that because of forgetfulness, changing reward and context at time τ only affects the private estimate $\hat{\theta}_{\ell+1}$ as it only depends on the data from the previous episodes. We now prove this case by case

1. Case-I: For a fixed action sequence $k_{-\tau} = \{k_{i,1}, \dots, k_{i,\tau-1}, k_{i,\tau+1}, \dots, k_{i,T}\}_{i \in [M]} \cup \{k_{j,\tau}\}_{j \in [M], j \neq m}$ for some $m \in [M]$. We have $\mathcal{H}_{\ell-1} = \mathcal{H}'_{\ell-1}$. Fixing a time t in episode ℓ , by the nature of Algorithm 1, the randomness in \mathcal{M}_t only comes through the estimator $\hat{\theta}_\ell$. Let the distribution function of $\hat{\theta}_\ell$ on dataset \mathcal{D} be F and those on \mathcal{D}' be F' . For a fixed action sequence, since $\mathcal{H}_{\ell-1} = \mathcal{H}'_{\ell-1}$, these distribution functions are identical. Define the set: $\mathcal{O} = \{\hat{\theta} : \mathcal{M}(\hat{\theta}_t, C_t) = k_t\}$, then we have the following for $\ell \leq \ell_0$:

$$\begin{aligned} \mathbb{P}(\hat{k}_t = k_t; t_\ell \leq t \leq t_{\ell+1}) &= \int_{\mathbb{R}^d} \mathbf{1}(t_\ell \leq t \leq t_{\ell+1}; t \neq \tau) dF_{\hat{\theta}}(z) \\ &= \mathbb{P}(\hat{\theta}_t(\mathcal{H}_{t-1}) \in \cap_{t_\ell \leq t \leq t_{\ell+1}; t \neq \tau} \mathcal{O}_t) \\ &= \mathbb{P}(\hat{\theta}_t(\mathcal{H}'_{t-1}) \in \cap_{t_\ell \leq t \leq t_{\ell+1}; t \neq \tau} \mathcal{O}_t) \end{aligned}$$

2. Case-II ($\ell > \ell_0 + 1$): This argument follows from the previous case and using the Forgetfulness property, we get:

$$\mathbb{P}(\hat{k}_t(\mathcal{H}_t) = k_t; t_\ell \leq t \leq t_{\ell+1}) = \mathbb{P}(\hat{k}_t(\mathcal{H}'_t) = k_t; t_\ell \leq t \leq t_{\ell+1})$$

3. Case-III ($\ell = \ell_0 + 1$): Noting that \mathcal{H}_{t-1} and \mathcal{H}'_{t-1} are neighbouring datasets and that $\hat{\theta}_{t-1}$ is an (ϵ, δ) -DP estimator. Hence,

$$\begin{aligned} \mathbb{P}(\hat{k}_t(\mathcal{H}) = k_t; t_\ell \leq t \leq t_{\ell+1}) &= \int_{\mathbb{R}^d} \mathbf{1}(t_{\ell_0+1} \leq t \leq t_{\ell_0+2}; t \neq \tau) dF(\hat{\theta}_\ell)(z) \\ &= \mathbb{P}(\hat{\theta}_t(\mathcal{H}_{\ell_0}) \in \cap_{t_{\ell_0} \leq t \leq t_{\ell_0+2}; t \neq \tau} \mathcal{O}_t) \\ &= \exp(\epsilon) \cdot \mathbb{P}(\hat{\theta}_t(\mathcal{H}'_{\ell_0}) \in \cap_{t_{\ell_0} \leq t \leq t_{\ell_0+2}; t \neq \tau} \mathcal{O}_t) + \delta \\ &= \exp(\epsilon) \cdot \mathbb{P}(\hat{k}_t = k_t; t_\ell \leq t \leq t_{\ell+1}; t \neq \tau) + \delta \end{aligned}$$

Combining these cases, we have:

$$\begin{aligned} \mathbb{P}(\mathcal{M}_{-\tau}(\mathcal{D}) = a_{-\tau}) &= \prod_{\ell \geq 0} \mathbb{P}(\hat{k}_t = k_t; t_\ell \leq t \leq t_{\ell+1}; t \neq \tau) \\ &= \exp(\epsilon) \cdot \prod_{\ell \geq 0} \mathbb{P}(\hat{k}_t = k_t; t_\ell \leq t \leq t_{\ell+1}; t \neq \tau) \\ &\quad + \delta \prod_{\ell \neq \ell_0+1} \mathbb{P}(\hat{k}_t(\mathcal{H}_t) \neq k_t; t_\ell \leq t \leq t_{\ell+1}; t \neq \tau) \\ &= \exp(\epsilon) \cdot \mathbb{P}(\mathcal{A}_{-\tau} = \mathcal{M}_{-\tau}) + \delta \end{aligned}$$

The proof for privacy in the central setting can be obtained by setting $M = 1$. □

901 **F Proof of Theorem 3**

902 *Proof.* Here, $\mathcal{G} \triangleq \cap_{\ell} \left\{ \max_b \mathbf{W}_b \leq K \frac{s^2 \log^2 d (\log \frac{B_{\ell}^*}{\delta})}{N_{\ell}^2 \epsilon^2} \right\}$, which holds with probability $1 - \delta$.

903 The regret decomposition is given by:

$$\begin{aligned}
\mathbb{E}[R(T)] &= \sum_{s=1}^t \mathbb{E} \left[X_{t, k_t^*}^{\top} \theta^* - X_{t, k_t}^{\top} \theta^* \right] \\
&= 2X_{\max} \|\theta^*\|_1 \sum_{1 \leq \ell \leq L} t_{\ell} + \sum_{s=t_L+1}^T \mathbb{E} \left[\left(X_{t, \theta^*}^{\top} \theta^* - X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} \right) + \left(X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} \right) \right. \\
&\quad \left. + \left(X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \theta^* \right) \right] \\
&= 2X_{\max} b_{\max} (2^L - 1) + \sum_{\ell \geq L} \sum_{t_{\ell} \leq t_{\ell+1}} \mathbb{E} \left[\left(X_{t, k_t^*}^{\top} \theta^* - X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} \right) \mathbf{1}_{\mathcal{G}} + \left(X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \theta^* \right) \mathbf{1}_{\mathcal{G}} \right] \\
&\quad + X_{\max} b_{\max} \mathbb{P}(\mathbf{1}_{\bar{\mathcal{G}}}) \\
&\leq 2X_{\max} b_{\max} (2^L - 1) + 2X_{\max} \sum_{\ell \geq L} \sum_{t_{\ell} \leq t_{\ell+1}} \mathbb{E} \left[\|\theta^* - \hat{\theta}_{\ell(t)}\|_1 \mathbf{1}_{\mathcal{G}} \right] + X_{\max} b_{\max} \mathbb{P}(\mathbf{1}_{\bar{\mathcal{G}}}) \\
&\leq 2X_{\max} b_{\max} 2^L + 2X_{\max} \sigma^2 \sqrt{s} \sum_{\ell \geq L} (t_{\ell+1} - t_{\ell}) \sqrt{K' \frac{(s^*)^2 \log^2 d}{N_{\ell}}} \\
&\quad + 2X_{\max} \sigma^2 \sqrt{s} \sum_{\ell \geq L} (t_{\ell+1} - t_{\ell}) \sqrt{K' \frac{s^2 \log^2 d \log N_{\ell}^2}{\epsilon^2} \log \left(\frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{DN_{\ell}} \right)} \\
&\leq 2X_{\max} b_{\max} \log T + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \sum_{\ell \geq L} \sqrt{K' \cdot N_{\ell}} \\
&\quad + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \log \left(\frac{s^2 \log^2 d \log \frac{\log \log T}{\delta}}{D} \right) \sqrt{\frac{K'}{\epsilon^2}} \\
&\leq 2X_{\max} b_{\max} \log T + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \sqrt{K' T} \\
&\quad + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \log \left(\frac{s^2 \log^2 d \log \frac{\log \log T}{\delta}}{D} \right) \sqrt{\frac{K'}{\epsilon^2}}
\end{aligned}$$

904 The per-step regret is given by:

$$\begin{aligned}
\Delta_{k_t}(t) &= x_{k_t^*}^{\top} \theta^* - x_{k_t}^{\top} \theta^* \\
&\leq x_{k_t^*}^{\top} \theta^* - x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} + x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} - x_{k_t}^{\top} \hat{\theta}_{\ell(t)} + x_{k_t}^{\top} \hat{\theta}_{\ell(t)} - x_{k_t}^{\top} \theta^* \\
&\leq \|x_{k_t^*}\|_{\infty} \|\theta^* - \hat{\theta}_{\ell(t)}\|_1 + \|x_{k_t}\|_{\infty} \|\theta^* - \hat{\theta}_{\ell(t)}\|_1 \\
&\leq 4\sigma X_{\max} s \sqrt{T \log d} + 4\sigma X_{\max} \sqrt{\frac{s^3 (\log d)^2 \log \left(\frac{1}{\delta} \right) \log^6 T}{\epsilon^2}}
\end{aligned}$$

905 where, the last-inequality follows from the previous lemma. Consider the event:

$$\mathcal{H}_t \triangleq \{x_{k_t^*}^{\top} \theta^* > \max_{k \neq k_t^*} x_k^{\top} \theta^* + p_t\}$$

906 Under $\mathcal{M}_{\ell} \cap \mathcal{A}_{\ell(t)}$, we have the following for any $k \neq k_t$:

$$\begin{aligned}
x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} - x_k^{\top} \hat{\theta}_{\ell(t)} &= \langle x_{k_t^*}, \hat{\theta} - \theta^* \rangle + \langle x_{k_t^*} - x_{k_t}, \theta^* \rangle + \langle x_{k_t}, \theta^* - \hat{\theta}_{\ell(t)} \rangle \\
&\geq -g_t + p_{t-1} - g_t
\end{aligned}$$

907 Choosing $p_t = 3g_{\ell(t)}$, with $g_{\ell(t)} = \sqrt{K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{N_{\ell}}} +$
 908 $\sqrt{K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{N_{\ell}^2 \epsilon^2}} \log \left(K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{DN_{\ell}} \right)$, we see that:

909 1. When $\alpha = 0$, we get the previously proven bound.

910 2. When $\alpha \in (0, 1)$:

$$\begin{aligned} \mathbb{E}(R(T)) &\leq 2X_{\max} b_{\max} (2^L - 1) \\ &\quad + 24X_{\max} \sigma^2 \sum_{\ell \geq L} \Delta_*^{-\alpha} N_{\ell-1} \left[\left(\frac{s^2 \log d}{N_{\ell-1}} \right)^{\frac{1+\alpha}{2}} + \left(\frac{s^3 (\log d)^2 \log(\frac{1}{\delta}) \log \frac{\log N_{\ell-1}}{\delta}}{N_{\ell-1}^2 \epsilon^2} \right)^{\frac{1+\alpha}{2}} \right] \end{aligned}$$

911 This implies that:

$$I_{\alpha} \leq \frac{s^{3(1+\alpha)} (\log d)^{\frac{(1+\alpha)}{2}}}{\Delta_*^{\alpha}} \left(\frac{T^{\frac{1-\alpha}{2}} - 1}{1 - \alpha} \right) + \Psi_{\alpha} \Delta_*^{-\alpha} \epsilon^{-(1+\alpha)} \{\log(1/\delta)\}^{\frac{(1+\alpha)}{2}} (s^3 \log^2 d)^{\frac{1+\alpha}{2}}$$

912 where, $\Psi_{\alpha} = \frac{1-T^{-\alpha}}{1-2^{-\alpha}} (\log T)^{2\alpha+2}$.

913 3. $\alpha = 1$: The previous bounds implies:

$$I_1 \leq \frac{s^6 (\log d)}{\Delta_*} \log T + \frac{\Psi_1}{\Delta_* \epsilon^2} s^3 \log d \log \left(\frac{1}{\delta} \right)$$

914 4. $\alpha > 1$: In this case, we get:

$$I_{\alpha} \leq \frac{s^{3(1+\alpha)} (\log d)^{\frac{(1+\alpha)}{2}}}{\Delta_*^{\alpha}} \left(\frac{1 - T^{-\frac{(\alpha-1)}{2}}}{\alpha - 1} \right) + \Psi_{\alpha} \Delta_*^{-\alpha} \epsilon^{-(1+\alpha)} \left(\log \left(\frac{1}{\delta} \right) \right)^{\frac{(1+\alpha)}{2}} (s^3 \log^2 d)^{\frac{(1+\alpha)}{2}}$$

915

□

916 **G Proof of Theorem 4**

917 *Proof.* Here, $\mathcal{G} \triangleq \cap_{\ell} \left\{ \max_b \mathbf{W}_b \leq K \frac{s^2 \log^2 d (\log \frac{B_{\ell}}{\delta})}{N_{\ell}^2 \epsilon^2} \right\}$, which holds with probability $1 - \delta$.

918 The regret decomposition is given by:

$$\begin{aligned}
\mathbb{E}[R(T)] &= \sum_{s=1}^t \mathbb{E} \left[X_{t, k_t^*}^{\top} \theta^* - X_{t, k_t}^{\top} \theta^* \right] \\
&= 2X_{\max} \|\theta^*\|_1 \sum_{1 \leq \ell \leq L} N_{\ell} + \sum_{s=t_L}^T \mathbb{E} \left[\left(X_{t, \theta^*}^{\top} \theta^* - X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} \right) + \left(X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} \right) + \left(X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \theta^* \right) \right] \\
&= 2X_{\max} b_{\max} (2^L - 1) + \sum_{m \in [M]} \sum_{\ell \geq L} \sum_{t_{\ell} \leq t_{\ell+1}} \mathbb{E} \left[\left(X_{t, k_t^*}^{\top} \theta^* - X_{t, k_t^*}^{\top} \hat{\theta}_{\ell(t)} \right) \mathbf{1}_{\mathcal{G}} + \left(X_{t, k_t}^{\top} \hat{\theta}_{\ell(t)} - X_{t, k_t}^{\top} \theta^* \right) \mathbf{1}_{\bar{\mathcal{G}}} \right] + X_{\max} \\
&\leq 2X_{\max} b_{\max} (2^L - 1) + 2X_{\max} \sum_{m \in [M]} \sum_{\ell \geq L} \sum_{t_{\ell} \leq t_{\ell+1}} \mathbb{E} \left[\|\theta^* - \hat{\theta}_{\ell(t)}\|_1 \mathbf{1}_{\mathcal{G}} \right] + X_{\max} b_{\max} \mathbb{P}(\mathbf{1}_{\bar{\mathcal{G}}}) \\
&\leq 2X_{\max} b_{\max} (2^L - 1) + 2X_{\max} \sigma^2 \sqrt{s} \sum_{m \in [M]} \sum_{\ell \geq L} (t_{\ell+1} - t_{\ell}) \sqrt{K' \frac{(s^*)^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{M N_{\ell}}} \\
&+ 2X_{\max} \sigma^2 \sqrt{s} \sum_{m \in [M]} \sum_{\ell \geq L} (t_{\ell+1} - t_{\ell}) \sqrt{K' \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{M^2 N_{\ell}^2 \epsilon^2}} \log \left(\frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{D N_{\ell}} \right) \\
&\leq 2X_{\max} b_{\max} \log T + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \sum_{\ell \geq L} \sqrt{K' \cdot M N_{\ell}} \\
&+ 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \log \left(\frac{s^2 \log^2 d \log \frac{\log \log T}{\delta}}{D} \right) \sqrt{\frac{K'}{\epsilon^2}} \\
&\leq 2X_{\max} b_{\max} \log T + 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \sqrt{K' M T} \\
&+ 2X_{\max} \sigma^2 s^{3/2} (\log d) \log \frac{\log \log T}{\delta} \log \left(\frac{s^2 \log^2 d \log \frac{\log \log T}{\delta}}{D} \right) \sqrt{\frac{K'}{\epsilon^2}}
\end{aligned}$$

919 The per-step regret is given by:

$$\begin{aligned}
\Delta_{k_t}(t) &= x_{k_t^*}^{\top} \theta^* - x_{k_t}^{\top} \theta^* \\
&\leq x_{k_t^*}^{\top} \theta^* - x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} + x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} - x_{k_t}^{\top} \hat{\theta}_{\ell(t)} + x_{k_t}^{\top} \hat{\theta}_{\ell(t)} - x_{k_t}^{\top} \theta^* \\
&\leq \|x_{k_t^*}\|_{\infty} \|\theta^* - \hat{\theta}_{\ell(t)}\|_1 + \|x_{k_t}\|_{\infty} \|\theta^* - \hat{\theta}_{\ell(t)}\|_1 \\
&\leq 4\sigma X_{\max} s \sqrt{M T \log d} + 4\sigma X_{\max} \sqrt{\frac{s^3 (\log d)^2 \log \left(\frac{1}{\delta} \right) \log^6 T}{\epsilon^2}}
\end{aligned}$$

920 where, the last-inequality follows from the previous lemma. Consider the event:

$$\mathcal{H}_t \triangleq \{x_{k_t^*}^{\top} \theta^* > \max_{k \neq k_t^*} x_k^{\top} \theta^* + p_t\}$$

921 Under $\mathcal{M}_{\ell} \cap \mathcal{A}_{\ell(t)}$, we have the following for any $k \neq k_t$:

$$\begin{aligned}
x_{k_t^*}^{\top} \hat{\theta}_{\ell(t)} - x_k^{\top} \hat{\theta}_{\ell(t)} &= \langle x_{k_t^*}, \hat{\theta} - \theta^* \rangle + \langle x_{k_t^*} - x_{k_t}, \theta^* \rangle + \langle x_{k_t}, \theta^* - \hat{\theta}_{\ell(t)} \rangle \\
&\geq -g_t + p_{t-1} - g_t
\end{aligned}$$

922 Choosing $p_t = 3g_{\ell(t)}$, with $g_{\ell(t)} = \sqrt{K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{N_{\ell}}} +$

923 $\sqrt{K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{N_{\ell}^2 \epsilon^2}} \log \left(K' \cdot \frac{s^2 \log^2 d \log \frac{\log N_{\ell}}{\delta}}{D N_{\ell}} \right)$, we see that:

924 1. When $\alpha = 0$, we get the previously proven bound.

925 2. When $\alpha \in (0, 1)$:

$$\mathbb{E}(R(T)) \leq 2X_{\max}b_{\max}(2^L-1)+24X_{\max}\sigma^2 \sum_{\ell \geq L} \Delta_*^{-\alpha} N_{\ell-1} \left[\left(\frac{s^2 \log d}{N_{\ell-1}} \right)^{\frac{1+\alpha}{2}} + \left(\frac{s^3 (\log d)^2 \log(\frac{1}{\delta}) \log \frac{\log N_{\ell-1}}{N_{\ell-1}^2 \epsilon^2}}{N_{\ell-1}^2 \epsilon^2} \right)^{\frac{1+\alpha}{2}} \right]$$

926 This implies that:

$$I_\alpha \leq \frac{s^{3(1+\alpha)} (\log d)^{\frac{(1+\alpha)}{2}}}{\Delta_*^\alpha} \left(\frac{(MT)^{\frac{1-\alpha}{2}} - 1}{1-\alpha} \right) + \Psi_\alpha \Delta_*^{-\alpha} \epsilon^{-(1+\alpha)} \{\log(1/\delta)\}^{\frac{(1+\alpha)}{2}} (s^3 \log^2 d)^{\frac{1+\alpha}{2}}$$

927 where, $\Psi_\alpha = \frac{1-(MT)^{-\alpha}}{1-2^{-\alpha}} (\log MT)^{2\alpha+2}$.

928 3. $\alpha = 1$: The previous bounds implies:

$$I_1 \leq \frac{s^6 (\log d)}{\Delta_*} \log T + \frac{\Psi_1}{\Delta_* \epsilon^2} s^3 \log d \log\left(\frac{1}{\delta}\right)$$

929 4. $\alpha > 1$: In this case, we get:

$$I_\alpha \leq \frac{s^{3(1+\alpha)} (\log d)^{\frac{(1+\alpha)}{2}}}{\Delta_*^\alpha} \left(\frac{1 - (MT)^{\frac{-(\alpha-1)}{2}}}{\alpha-1} \right) + \Psi_\alpha \Delta_*^{-\alpha} \epsilon^{-(1+\alpha)} \left(\log\left(\frac{1}{\delta}\right) \right)^{\frac{(1+\alpha)}{2}} (s^3 \log^2 d)^{\frac{(1+\alpha)}{2}}$$

930

□

931 H Proof of Theorem 5

932 **Definition 6** (Collinearly-dependent policies (Huang et al., 2021)). *Two clients i and j are called*
 933 *collinear if there exist an arm $k \in [K]$ and a subset $S \subset [M]$ such that the following conditions*
 934 *are satisfied: 1) $X_{i,k} \in \text{span}(\{X_{m,k} | m \in S\})$ and 2) $x_{i,k} \in \text{span}(\{X_{m,k} | m \in S\} \cup X_{j,k})$. A*
 935 *collinearly-dependent policy if for any two clients i and j that are not collinear, if the action of client*
 936 *is independent of the action of j under a policy π .*

937 Let $\mathcal{F}_{X,r}(\theta)$ denote the space of joint distributions of contexts and rewards upto time T , when the
 938 underlying parameter is θ . In particular,

$$\mathcal{F}_{X,r} = \otimes_{t \in [T]} \mathcal{F}_{X_t, r_t}$$

939 where, \mathcal{F}_{x_t, r_t} denotes the joint distribution of contexts $\{x_1(t), x_2(t), \dots, x_K(t)\} \in \mathbb{R}^{dK}$ and rewards
 940 $\{r_1(t), r_2(t), \dots, r_K(t)\} \in \mathbb{R}^d$ at time t induced through $\{x_1(t), \dots, x_K(t)\} \sim \mathbb{P}_X$ (where, the
 941 marginal for $x_i(t) \sim \mathbb{P}_i$ and the minimum and maximum eigenvalues of the marginal covariance
 942 are given by $\lambda_{\min}, \lambda_{\max}$). $r_i(t) = x_i(t)^\top \theta + \epsilon_i(t)$, $\epsilon_i(t) \sim \mathbb{P}_\epsilon$ independently for all ϵ , where \mathbb{P}_ϵ
 943 is sub-Gaussian with variance proxy σ^2 . For every choice of $\theta \in \mathbb{R}^d$, we get a bandit instance,
 944 $\mathbb{P}_{X,r} \in \mathcal{F}_{X,r}(\theta)$.

945 A policy π is a sequence of (randomized) maps $\pi_t : \mathcal{H}_{t-1}^\pi \times X_t \mapsto \Delta(K)$, where $\mathcal{H}_{t-1} =$
 946 $\{(x_{a_s}, r_{a_s}) : 1 \leq s \leq t\}$ denotes the history upto time t and $\Delta([K])$ denotes the probability
 947 simplex over the set of arms. The minmax regret in the private setting is defined as:

$$R_{(\epsilon, \delta)}^{\min \max}(T) = \inf_{\pi \in \Pi(\epsilon, \delta)} \sup_{P_{X,r} \in \mathcal{F}_{X,r}} \mathbb{E}[R(T)]$$

948 where, the expectation is taken wrt randomness due to interaction of the bandit instance $P_{X,r}$ and
 949 policy π ; the infimum is over all possible history-adapted (ϵ, δ) -JDP policies and the supremum is
 950 wrt all problem instances with sparsity s .

951 1. Constructing the hard instance:

952 We construct several hard instances as follows: consider the case of 2-armed problem with
 953 $\theta \in \tilde{\Theta} = \{\theta \in \mathbb{R}^d : \theta_i \in \{-r_{\min}, 0, r_{\max}\}, \|\theta\|_0 = s\}$, where s is the true sparsity parameter.
 954 Then, $|\tilde{\Theta}| = \binom{d}{s} 2^s$. Let

$$\delta(t) = \{\delta : \|\theta - \theta'\|_2 \geq \delta', \theta, \theta' \in \tilde{\Theta}, d_H(\theta, \theta') \geq t\}$$

955 where, d_H is the Hamming distance. Considering a t -packing of $\tilde{\Theta}$ in the Hamming distance, this
 956 leads to $\delta(t)$ packing of $\tilde{\Theta}$ in the ℓ_2 -distance Duchi and Wainwright (2013). With our construction,
 957 this leads to $\delta(t) > \max\{1, \sqrt{t}\}r$. Let $\Theta^* = \{\theta_1, \theta_2, \dots, \theta_M\}$ be the elements of this packing
 958 with $t >$ and $M > cs \log \frac{d}{s}$ for some absolute constant c . For $\theta, \theta' \in \tilde{\Theta}$, $\|\theta - \theta'\| \geq \alpha$.

959 2. Bounds between the reward-context distribution:

960 For $\theta, \theta' \in \tilde{\Theta}$, we bound the KL-divergence between the context and reward distribution, which is
 961 bounded as:

$$\begin{aligned} \text{KL}(P(\cdot|\theta) \| P(\cdot|\theta')) &= MKt \text{KL}(\mathcal{N}(x^\top \theta, \sigma^2), \mathcal{N}(x^\top \theta', \sigma^2)) \\ &= \frac{MKt \sigma_x^2 \|\theta - \theta'\|}{2\sigma^2} \\ &\leq \frac{MKt \sigma_X^2 s r^2}{\sigma^2} \end{aligned}$$

962 Using Pinsker's inequality, the Total variation distance can be bounded as:

$$\begin{aligned} \text{TV}(P(\cdot|\theta) \| P(\cdot|\theta')) &\leq MKt \sqrt{\frac{1}{2} \mathbb{E}[\text{KL}(\mathcal{N}(x^\top \theta, \sigma^2), \mathcal{N}(x^\top \theta', \sigma^2))]} \\ &\leq MKt \frac{\sigma^2}{\sigma_x^2} r \sqrt{s} \end{aligned}$$

3. Reduction to Estimation Problem

Given a policy π and its associated estimator $\hat{\theta}_t$ be the maximizer of

$$\mathbb{P}(k_t^\pi = 1, z_t^\top \theta \geq 0 | H_{t-1}^\pi) + \mathbb{P}(k_t^\pi = 2, z_t^\top \theta \leq 0 | H_{t-1}^\pi)$$

where, the max is over all possible random maps from history induced by π which are (ϵ, δ) -DP.

Let $\hat{\Theta}_t^{\epsilon, \delta}$ denote this space of random (ϵ, δ) -DP maps and π be an (ϵ, δ) -JDP policy. Therefore, from Theorem ??, we can conclude that:

$$\inf_{\pi \in \Pi} \sup_{P_{X,r} \in \mathcal{F}} \mathbb{E}[R^\pi(T)] \geq \inf_{\pi \in \Pi} \sup_{P_{X,r} \in \mathcal{F} - \text{construct}} \mathbb{E}[R^\pi(T)] = \Omega \left(\frac{1}{r\sqrt{s}} \sum_{t=1}^T \inf_{\hat{\theta}_t \in \hat{\Theta}_t^{\epsilon, \delta}} \mathbb{E}_\nu [\|\theta - \hat{\theta}_t\|_2^2] \right)$$

4. We divide this step into two parts, where for the ϵ -DP case we use Fano's inequality (Acharya et al., 2021) and the connection between ρ -zCDP and (ϵ, δ) -DP (Bun et al., 2018) and ρ -zCDP Fano's inequality (Kamath et al., 2022) to get the lower bound for any (ϵ, δ) -JDP.

(a) Analysis under ϵ -DP. For the non-private lower bound, we have:

$$\mathbb{E}[\|\theta_t - \theta^*\|_2^2] \geq \frac{\max\{1, s/4\}r^2}{2} \left(1 - \frac{2MKt\sigma_x^2 sr^2}{\sigma^2} + \log 2 \right) / (cs \log(d/s))$$

For the private lower-bound, we have:

$$\mathbb{E}[\|\theta_t - \theta^*\|_2^2] \geq 0.4 \max\{1, s/4\}r^2 \min \left\{ 1, \exp \left(cs \log(d/s) - 10\epsilon t MK \frac{\sigma_x}{\sigma} r \sqrt{s} \right) \right\}$$

For $s > 4$ and $\epsilon > 0$, we have:

$$\mathbb{E}[\|\theta_t - \theta^*\|_2^2] \geq \frac{sr^2}{4} \max \left\{ 0.5 \left(1 - \frac{2t\sigma_x^2 sr^2 MK}{cs \log(d/s)} \right), 0.4 \exp \left(cs \log(d/s) - 10\epsilon t \sqrt{s} r \frac{\sigma_x}{\sigma} MK \right) \right\}$$

Hence, for $K = 2$, the minimax-regret is given by:

$$\inf_{\pi \in \Pi} \sup_{P_{X,r} \in \mathcal{F}} \mathbb{E}[R(T)] \geq \frac{r\sqrt{s}MT}{4} \max \left\{ 1 - \frac{4MTsr^2\sigma_x^2}{\sigma^2} + \log 2, \exp \left(cs \log(d/s) - 20\epsilon MT \frac{\sigma_x}{\sigma} r \sqrt{s} \right) \right\}$$

Setting $r^2 = \frac{\sigma^2 \log(d/s)}{8MT}$, the non-private lower bound becomes:

$$\inf_{\pi \in \Pi} \sup_{P_{X,r} \in \mathcal{F}} \mathbb{E}[R(T)] \geq \frac{1}{8\sqrt{2}} \sigma \sqrt{MTs \log(d/s)} \left(1 - \frac{\sigma_x^2}{2c} - \frac{\log 2}{cs \log(d/s)} \right)$$

Noting that $\sigma_x^2 = (\log d)^{-1}$, $d > 9$ and $s \log(d/s) > 4c^{-1} \log 2$, our non-private lower bound is $\Omega \left(\sqrt{sMT \log(d/s)} \right)$.

For the private part, we set $r = \frac{c\sigma\sqrt{s \log(d/s)}}{20\epsilon\sigma_x MT}$, to get:

$$\inf_{\pi \in \Pi} \sup_{P_{X,r} \in \mathcal{F}} \mathbb{E}[R(T)] \geq \sigma \frac{s \log(d/s) \log^{1/2}(d)}{\epsilon} \geq \sigma \frac{\log^{3/2}(d/s)}{\epsilon}$$

Combining these results, we get:

$$R_\epsilon^{\text{minimax}}(T) = \Omega \left(\max \{ s \log^{3/2}(d/s) \epsilon^{-1}, \sqrt{sMT \log(d/s)} \} \right)$$

(b) For (ϵ, δ) -JDP, we construct $\rho(\epsilon, \delta)$ such that any (ϵ, δ) -DP algorithm is ρ -zCDP and thus $\Theta_t^{\epsilon, \delta} \subset \Theta_t^\rho$ where the latter is the space of all ρ -zCDP estimators constructed using the entire history H_{t-1} . Other arguments are as before.

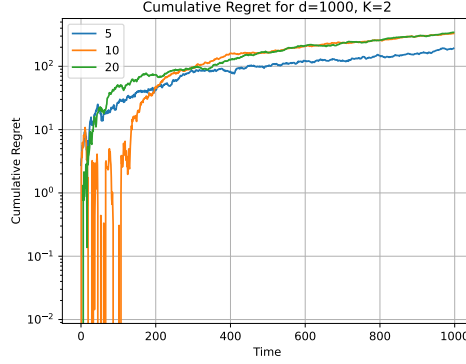


Figure 1: Regret in central setting. The synthetic instances are constructed as in Ariu et al. (2020) and the sparsity parameter is varied. We see that the regret behaves well across different values of sparsity.

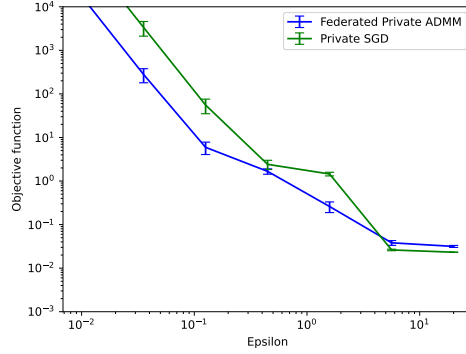


Figure 2: Decrement in estimation error for distributed LASSO.

I Experimental Evaluation

The experiments are done on a MacBook Air with an Apple M1 chip, 16 GB memory and 10 core CPU. All codes are written in Python3 using several open source packages. The running time for all experiments ranges from less than a minute to a few hours. The code is available at: [this link](#).

I.1 Simulations with non-bandit feedback

We illustrate the performance of Algorithm 3 on a distributed LASSO problem, i.e., solving:

$$\min_x \frac{1}{2n} \|Ax - b\|^2 + \kappa \|x\|_1$$

where, $A \in \mathbb{R}^{n \times p}$ and $b \in \mathbb{R}^p$ is a vector of regression targets. We generate synthetic data by drawing A as random vectors from p -dimensional sphere and x from uniform distribution with support size 10. Labels are then obtained by taking $b = Ax + \eta$ where $\eta \sim \mathcal{N}(0, 0.01)$ with $n = 500$ and $p = 32$.

The benchmark comparison is performed with `scikit-learn` and we obtain the best parameter using cross-validation. We also compare our algorithm with ADMM with soft-thresholding (see (Cyffers et al., 2023)) and DP-SGD where noise is added to the gradients. For both approaches, we tune step-size and clipping threshold using grid search. These parameters are tuned on the smallest privacy budget and used for all other values.

We report the objective function value on test set at the end of training and convert them to (ϵ, δ) -DP for the sake of comparison with $\delta = 1e - 6$. Each user is randomly sampled with probability 10%.

999 J Auxiliary Results

1000 In this appendix, we provide a collection of existing results used in our proofs.

1001 **Theorem 6** (Fixed-point iteration Combettes and Pesquet (2021)). *Assume that R is a τ -contractive*
 1002 *operator with fixed point u_* for $\tau \in [0, 1)$. Then there exists a learning rate $\lambda_b \in (0, 1]$ such that the*
 1003 *iterates of Algorithm 1 satisfy:*

$$\mathbb{E} [\|u_{k+1} - u^*\|^2 | \mathcal{F}_0] \leq \left(1 - \frac{(1 - \tau)}{8}\right)^B D$$

1004 where $D = \max \|u_0 - u^*\|$ is the diameter of the domain, p is the dimension of u , $\sigma^2 > 1 - \tau$ is the
 1005 variance of variance of gaussian noise and $\mathbb{E}[\|e_k\|^2] \leq \xi^2$ for some $\xi \geq 0$.

1006 **Theorem 7** (Laplace concentration Cai et al. (2020)). *Consider $w \in \mathbb{R}^k$ with $w_1, w_2, \dots, w \stackrel{iid}{\sim}$*
 1007 *Laplace(λ). For every $C > 1$, we have:*

$$\begin{aligned} \mathbb{P}(\|w\|_2^2 > kC^2\lambda^2) &\leq k \exp(-C) \\ \mathbb{P}(\|w\|_\infty > C^2\lambda^2 \log^2 k) &\leq \exp(-(C - 1) \log k) \end{aligned}$$

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