AN ONLINE LEARNING APPROACH TO PROMPT-BASED SELECTION OF GENERATIVE MODELS

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ABSTRACT

Selecting a sample generation scheme from multiple text-based generative models is typically addressed by choosing the model that maximizes an averaged evaluation score. However, this score-based selection overlooks the possibility that different models achieve the best generation performance for different types of text prompts. An online identification of the best generation model for various input prompts can reduce the costs associated with querying sub-optimal models. In this work, we explore the possibility of varying rankings of text-based generative models for different text prompts and propose an online learning framework to predict the best data generation model for a given input prompt. The proposed framework adapts the kernelized contextual bandit (CB) methodology to a CB setting with shared context variables across arms, utilizing the generated data to update a kernel-based function that predicts which model will achieve the highest score for unseen text prompts. Additionally, we apply random Fourier features (RFF) to the kernelized CB algorithm to accelerate the online learning process and establish a $\mathcal{O}(\sqrt{T})$ regret bound for the proposed RFF-based CB algorithm over T iterations. Our numerical experiments on real and simulated text-to-image and image-to-text generative models show RFF-UCB performs successfully in identifying the best generation model across different sample types.

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1 INTRODUCTION

031 Text-based generative artificial intelligence (AI) has found numerous applications in various engi-032 neering tasks. A prompt-based generative AI represents a conditional generative model that pro-033 duces samples given an input text prompt. Over the past few years, several frameworks using diffu-034 sion models and generative adversarial networks have been proposed to perform text-guided sample generation tasks for various data domains including image, audio, and video (Reed et al., 2016; Pan et al., 2018; Xu et al., 2018; Ding et al., 2021; Singer et al., 2022; Huang et al., 2023; Podell et al., 2024). The multiplicity of developed prompt-based models has led to significant interest in 037 developing evaluation mechanisms to rank the existing models and find the best generation scheme. To address this task, several evaluation metrics have been proposed to quantify the fidelity and relevance of samples created by prompt-based generative models, such as CLIPScore (Hessel et al., 040 2021) and PickScore (Kirstain et al., 2023). 041

The existing model selection methodologies commonly aim to identify the generative model with 042 the highest relevance score, producing samples that correlate the most with input text prompts. A 043 well-known example is the CLIPScore for image generation models, measuring the expected align-044 ment between the input text and output image of the model using the CLIP embedding (Radford 045 et al., 2021b). While the best-model identification strategy has been frequently utilized in gener-046 ative AI applications, this approach does not consider the possibility that the involved models can 047 perform differently across text prompts. However, it is possible that one model outperforms another 048 model in responding to text prompts from certain categories, while that model performs worse in generating samples for other text categories. Figure 1 shows one such example where two standard text-to-image models exhibit different rankings on text prompts with the terms "dog" and "car". In 051 general, the different training sets and model architectures utilized to train text-based models can result in the models' varying performance in response to different text prompts, which is an important 052 consideration in the selection of text-based generative models for various text prompts. We provide more examples in Figures 17 and 18 in the Appendix.

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Figure 1: Prompt-based generated images from Stable Diffusion and PixArt- α : Stable Diffusion attains a higher CLIPScore in generating type-2 prompts (36.10 versus 35.68) while underperforms for type-1 prompts (36.37 versus 37.24).

In this work, we aim to develop a learning algorithm to identify the best generative model for a given 082 input prompt, using observed prompt/generated samples collected from the models in the previous 083 sample generation iterations. Since the goal of such text-based model selection is to minimize 084 the data queries from suboptimal generative models for an input text prompt, we view the model 085 selection task as an online learning problem, where after each data generation the learner updates a function predicting which generative model performs the best in response to different text prompts. 087 Here, the goal of the online learner is to utilize the previously generated samples to accurately guess 880 the generation model with the best performance for the incoming text prompt. An optimal online 089 model selection method will result in a bounded regret value, measured in comparsion to the sample generation from the groundtruth-best model for the text prompts.

091 The described online learning task can be viewed as a *contextual bandit (CB)* problem studied in the 092 multi-armed bandit literature (Langford & Zhang, 2007; Li et al., 2010). In a CB task, the online 093 learner observes the context variable (the text prompt in our setting) and guesses the best arm for the 094 current input context. Specifically, we focus on the kernelised upper confidence bound (kernel-UCB) 095 approach and adapt this methodology to propose the Shared-Context Kernel UCB (SCK-UCB) for 096 an online prompt-based selection of generative models. According to the SCK-UCB approach, the learner utilizes a UCB-score from a kernel-based prediction function to choose the generative model for the incoming text prompt and subsequently update the kernel-based prediction rule based on 098 the generated data for the upcoming iterations. We prove that the proposed SCK-UCB achieves an 099 $O(\sqrt{T})$ regret bound over a horizon of T iterations. 100

Since the user applying the CB-based model selection approach may have limited compute power and not be able to afford growing computational costs in the online learning process, we propose to utilize the random Fourier features (RFF) framework (Rahimi & Recht, 2007b) to balance the computational load between the iterations of SCK-UCB. We discuss that in the kernel-UCB methods, including our proposed SCK-UCB, the computational cost per iteration will grow cubically as $O(t^3)$ with iteration t. To address the growing cost per iteration of kernel-UCB, we leverage the RFF approach and develop the proxy RFF-UCB algorithm which approximates the solution to SCK-UCB, while the computational costs grow only linearly O(t). We show that the regret bound for SCK- UCB will approximately hold for RFF-UCB, and therefore, RFF-UCB provides an efficient proxy to the SCK-UCB algorithm, which can be run in devices with lower computation budget.

Finally, we present the results of several numerical experiments to show the efficacy of our pro-111 posed SCK-UCB and RFF-UCB in the online selection of conditional generative models. In our ex-112 periments, we simulate several text-to-image and image-captioning (image-to-text) models, where 113 different models lead to different rankings of CLIPScore values across sample types. Our numeri-114 cal results suggest a fast convergence of the proposed online learning algorithms to the best model 115 available for different prompt types. Moreover, we apply the RFF-UCB method to several stan-116 dard text-to-image models, and show how the algorithm can infer the model with higher CLIPScore 117 with a growing accuracy as the iteration grows. In our experiments, the proposed SCK-UCB and 118 RFF-UCB outperform the greedy baselines without any bonus term to encourage exploration in the learning process. The main contributions of our work can be summarized as: 119

- Studying the prompt-based selection of conditional generative models to improve the performance scores over every individual model
- Developing the contextual bandit-based SCK-UCB and RFF-UCB algorithms for the online selection of prompt-based generative models
- Providing the theoretical analysis of the regret and computational costs of SCK-UCB and RFF-UCB online learning methods
- Presenting numerical results on the online selection of generative models based on the incoming prompt using SCK-UCB and RFF-UCB
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2 RELATED WORK

133 (Automatic) Evaluation of conditional generative models. Evaluating the conditional genera-134 tive models has been studied extensively in the literature. For text-to-image (T2I) generation, earlier 135 methods primarily rely on the Inception score (Salimans et al., 2016) and Fréchet inception dis-136 tance (Heusel et al., 2017). More recent works propose reference-free metrics for robust automatic 137 evaluation of T2I and image captioning, with notable examples being CLIPScore (Hessel et al., 138 2021) and PickScore (Kirstain et al., 2023). Kim et al. (2022) propose a mutual-information-based 139 metric, which attains consistency across benchmarks, sample parsimony, and robustness. To pro-140 vide a holistic evaluation of T2I models, several works focus on multi-objective evaluation. Astolfi 141 et al. (2024) propose to evaluate conditional image generation in terms of prompt-sample consis-142 tency, sample diversity, and fidelity. Kannen et al. (2024) introduce a framework to evaluate T2I models regarding cultural awareness and cultural diversity. Masrourisaadat et al. (2024) examine 143 the performance of several T2I models in generating images such as human faces and groups and 144 present a social bias analysis. Another line of study explores evaluation approaches using large 145 language models (LLMs). Tan et al. (2024) develop LLM-based evaluation protocols that focus on 146 the faithfulness and text-image alignment. Peng et al. (2024) introduce a GPT-based benchmark for 147 evaluating personalized image generation. For evaluation of text-to-video (T2V) generation, Huang 148 et al. (2024) introduce VBench as a comprehensive evaluation of T2V models in terms of quality 149 and *consistencey*. 150

(Kernelized) Contextual bandits. The contextual bandits (CB) is an efficient framework for 151 online decision-making with context information (Langford & Zhang, 2007; Foster et al., 2018), 152 which is widely adopted in domains such as recommendation system and online advertisement (Li 153 et al., 2010). A key to its formulation is the relationship between the context (vector) and the ex-154 pected reward. In linear CB, the reward is assumed to be linear to the context vector (Li et al., 2010; 155 Chu et al., 2011). To incorporate non-linearity, Valko et al. (2013) propose kernelized CB, which 156 assumes the rewards are linear-realizable in a reproducing kernel Hilbert space (RKHS). However, 157 the proposed algorithm requires solving a kernel ridge regression per iteration, whose computation 158 and required space have polynomial dependence on the number of iterations. To address this prob-159 lem, a line of study leverages the assumption that the kernel matrix is often approximately low-rank and uses Nyström approximations (Calandriello et al., 2019; 2020; Zenati et al., 2022). [Recently, 160 a line of study utilizes (contextual) bandit algorithms to improve the performance of generative 161 models (Chen et al., 2024; Lin et al., 2024).]

¹⁶² 3 PRELIMINARIES

3.1 CLIPSCORE

166 CLIPScore (Hessel et al., 2021) is a widely used automatic metric to evaluate the alignment of text-167 to-image/video (T2I/V) and image captioning models. Let $(y, x) \in \mathcal{Y} \times \mathcal{X}$ be any *text-image pair*. 168 We denote by $c_y \in \mathbb{S}^{d-1} := \{z \in \mathbb{R}^d : ||z||_2 = 1\}$ and $v_x \in \mathbb{S}^{d-1}$ the (normalized) embeddings 169 of text $y \in \mathcal{Y}$ and image $x \in \mathcal{X}$, respectively, both extracted by CLIP (Radford et al., 2021a). The 170 CLIPScore (Hessel et al., 2021) is given by

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and note that $\cos(v_x, c_y) = \langle v_x, c_y \rangle$ as we operate under the normalized embeddings. Further, for a video $X := \{x^{(l)}\}_{l=1}^{L}$ consisting of L frames, where $x^{(l)}$ is the *l*-th frame, the score is the averaged frame CLIPScore, that is,

 $\mathbf{CLIPScore}^{\mathsf{T2I}}(y, x) := \max\{0, 100 \cdot \cos(\boldsymbol{v}_x, \boldsymbol{c}_y)\},\$

$$\text{CLIPScore}^{\text{T2V}}(y, X) := \frac{1}{L} \sum_{l=1}^{L} \text{CLIPScore}^{\text{T2I}}(y, x^{(l)}).$$
(2)

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3.2 KERNEL METHODS AND RANDOM FOURIER FEATURES

Let $\phi : \mathbb{R}^d \to \mathcal{H}$ denote a mapping from the *primal space* \mathbb{R}^d to the (possibly infinite-183 dimensional) associated reproducing kernel Hilbert space (RKHS) H. The corresponding kernel 184 function is defined by $k(y,y') := (\phi(y))^\top \phi(y')$ for any $y, y' \in \mathbb{R}^d$, where we use matrix notation 185 $h_1^{\top}h_2 := \langle h_1, h_2 \rangle_{\mathcal{H}}$ to denote the inner product of two elements $h_1, h_2 \in \mathcal{H}$. The kernel func-186 tion k is positive definite if $\sum_{i=1}^{n} \sum_{j=1}^{n} c_i c_j k(y_i, y_j) \ge 0$ for any $n \in \mathbb{N}_+, y_1, \cdots, y_n \in \mathbb{R}^d$, 187 and $c_1, \dots, c_n \in \mathbb{R}$. In other words, the kernel matrix $K := [k(y_i, y_j)]_{i,j=1}^n \in \mathbb{R}^{n \times n}$ is al-188 ways positive semi-definite (PSD). Further, a positive definite kernel function is shift invariant if 189 k(y, y') := k(y - y') for any $y, y' \in \mathbb{R}^d$. An example is the *radial basis function* (RBF) kernel, i.e., 190 $k_{\text{RBF}}(y, y') = \exp(-\|y - y'\|_2^2/(2\sigma^2))$ with $\sigma > 0$. 191

Kernel ridge regression (KRR). [Given empirical data $(y_1, s_1), \dots, (y_n, s_n)$, where $\{y_i \in \mathbb{R}^d\}_{i=1}^n$ are *dependent variables* and $\{s_i \in \mathbb{R}\}_{i=1}^n$ are *target variables*, respectively], the kernel method assumes the existence of $w^* \in \mathcal{H}$ such that $\mathbb{E}[s_i|y_i] = (\phi(y))^\top w^*$ for any $i = 1, \dots, n$. Let $\alpha \ge 0$ denote a *regularization parameter*. KRR constructs the estimator $\hat{s}_{\text{KRR}}(y) := k_y^\top (K + \alpha I_n)^{-1} v$ for any $y \in \mathbb{R}^d$, where $K = [k(y_i, y_j)]_{i,j=1}^n \in \mathbb{R}^{n \times n}$ is the kernel matrix, $v := [s_1, \dots, s_n]^\top \in \mathbb{R}^n$, and $k_y = [k(y_1, y), \dots, k(y_n, y)]^\top \in \mathbb{R}^n$. The estimator can be interpreted as first estimating w^* by ridge regression $\hat{w} := \arg \min_{w \in \mathcal{H}} \sum_{i=1}^n ((\phi(y_i))^\top w - s_i)^2 + \alpha \|w\|$, where $\|w\| := \sqrt{w^\top w}$ for any $w \in \mathcal{H}$, and then making the prediction $\hat{s}_{\text{KRR}}(y) = (\phi(y))^\top \hat{w}$.

200 **Random Fourier features (RFF).** One problem of KRR is that it scales poorly with the size n201 of the empirical data, i.e., computing the KRR estimator generally requires $\Omega(n^3)$ time and $\Omega(n^2)$ 202 memory. To address this problem, Rahimi & Recht (2007a) propose to scale up kernel methods by RFF sampling. Specifically, the Bochner's Theorem (Rudin, 2017) implies that for any (properly 203 scaled) shift-invariant kernel k(y, y') = k(y - y'), there exists a distribution $p \in \Delta(\mathbb{R}^{d})$ such that 204 $k(y, y') = \mathbb{E}_{w \sim p}[e^{iw^{\top}(y-y')}]$, where $e^{i\theta} := \cos \theta + i \cdot \sin \theta$ for any $\theta \in \mathbb{R}$ and *i* is the *imaginary unit*. Therefore, the idea of RFF is to sample $w_1, \cdots, w_s \sim p$ and approximate k(y, y') by the empirical 205 206 mean $s^{-1} \sum_{i=1}^{s} e^{iw_j^\top (y-y')}$ to within ϵ with only $s = O(d\epsilon^{-2} \log(1/\epsilon^2))$. Since the kernel $k(\cdot)$ is 207 208 real, we can replace the complex exponentials with cosines and define

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$$\varphi(y) := \sqrt{\frac{2}{s}} \cdot \left[\cos(w_j^\top y + b_j), \cdots, \cos(w_s^\top y + b_s)\right]^\top, \text{ where } w_j \stackrel{\text{i.i.d}}{\sim} p, b_j \stackrel{\text{i.i.d}}{\sim} \text{Unif}([0, 2\pi]) \quad (3)$$

and k(y, y') is approximated by $(\varphi(y))^{\top}\varphi(y')$. The resulting approximate KRR estimator $\widetilde{s}_{\text{KRR}}(y) := (\widetilde{\Phi}^* \widetilde{\Phi} + \alpha I_s)^{-1} \widetilde{\Phi}^* v$, where $\Phi := [\varphi(y_i)^{\top}]_{i=1}^n \in \mathbb{C}^{n \times s}$, can be computed in $O(ns^2)$ time and O(ns) memory, giving substantial computational savings if $s \ll n$ (Avron et al., 2017). For the RBF kernel, the distribution p_{RBF} is the multivariate Gaussian $\mathcal{N}(0, \sigma^{-2} \cdot I_d)$.

216 **PROMPT-BASED SELECTION AS CONTEXTUAL BANDITS** 4 217

218 In this section, we introduce the framework of online prompt-based selection of generative models, 219 which is given in Protocol 6. Let $[N] := \{1, \dots, N\}$ for any positive integer $N \in \mathbb{N}_+$. We denote 220 by $\mathcal{G} := [G]$ the set of (prompt-based) generative models. The evaluation proceeds in $T \in \mathbb{N}_+$ iterations. At any iteration $t \in [T]$, a prompt $y_t \in \mathcal{Y}$ is drawn from a fixed distribution $\rho \in \Delta(\mathcal{Y})$ on the prompt space $\mathcal{Y} \subseteq \mathbb{S}^{d-1}$, e.g., (the normalized embedding of) a picture in image captioning 221 222 or a paragraph in text-to-image/video generation. Based on prompt y_t (and previous observation sequence), an algorithm \mathcal{A} picks model $g_t \in \mathcal{G}$ and samples an answer $x_t \sim P_{q_t}(\cdot|y_t)$, where 224 $P_q(\cdot|y) \in \Delta(\mathcal{X})$ is the conditional distribution of answers generated from any model $g \in \mathcal{G}$. The 225 quality of answer x_t is given by $s(y_t, x_t)$, where $s: \mathcal{Y} \times \mathcal{X} \to [-1, 1]$ is the score function. The 226 algorithm \mathcal{A} aims to minimize the *regret* 227

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Regret(T) := $\sum_{t=1}^{T} (s_{\star}(y_t) - s_{g_t}(y_t)),$ (4)

where we denote by $s_g(y) := \mathbb{E}_{x_q \sim P_q(\cdot|y)}[s(y, x_g)]$ the expected score of any model $g \in \mathcal{G}$ and $s_{\star}(y) := \max_{q \in \mathcal{G}} s_q(y)$ the optimal expected score, both conditioned to prompt y. 232

Protocol 1 Online Prompt-based Selection of Generative Models

Require: total iterations $T \in \mathbb{N}_+$, set of generators $\mathcal{G} = [G]$, prompt distribution $\rho \in \Delta(\mathcal{Y})$, score function $s: \mathcal{Y} \times \mathcal{X} \to [-1, 1]$, algorithm $\mathcal{A}: (\mathcal{Y} \times \mathcal{G} \times \mathbb{R})^* \times \mathcal{Y} \to \Delta(\mathcal{G})$

Initialize: observation sequence $\mathcal{D} \leftarrow \emptyset$ 1: for iteration $t = 1, 2, \cdots, T$ do

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Prompt $y_t \sim \rho$ is revealed. 2:

Algorithm \mathcal{A} picks model $g_t \sim \mathcal{A}(\cdot | \mathcal{D}, y_t)$ and samples an answer $x_t \sim P_{g_t}(\cdot | y_t)$. 3:

4: Score $s_t \leftarrow s(y_t, x_t)$ is assigned.

5: Update observation sequence $\mathcal{D} \leftarrow \mathcal{D} \cup \{(y_t, g_t, s_t)\}$.

6: end for

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5 AN OPTIMISM-BASED APPROACH FOR PROMPT-BASED SELECTION

247 Under the setting of online prompt-based selection, a key challenge is to learn the relationship 248 between the prompt and the expected score for each model. In this paper, we assume the scores are 249 linear to the prompt vector in the reproducing kernel Hilbert space (RKHS) with model-dependent 250 weights.

251 **Assumption 1** (Realizability). There exists a mapping $\phi : \mathbb{R}^d \to \mathcal{H}$ and weight $w_a^{\star} \in \mathcal{H}$ such that 252 score $s_g(y) = \langle y, w_g^* \rangle_{\mathcal{H}}$ for any prompt vector $y \in \mathbb{R}^d$ and model $g \in \mathcal{G}$. Further, it holds that $\|w_g^*\| \leq 1$, and $k(y, y) \leq \kappa^2$ and $\|\phi(y)\| \leq 1$ for any $y \in \mathcal{Y}$, where $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ is the kernel 253 254 function of the mapping ϕ . 255

Remark 1 (Shared context with kernelized rewards). We note that Assumption 1 is slightly different 256 from the one made in kernelized bandits (Valko et al., 2013; Zenati et al., 2022), where a context 257 is observed per each arm and assumes the existence of a shared weight. [However, in the prompt-258 based generation setting, there is a single prompt at each iteration and responses can vary across 259 the models, which leads to our formulation of shared context and model-dependent weights.] 260

Remark 2. Our assumption of linear-realizable scores in RKHS is motivated by the following ob-261 servations. First, the relationship between the prompt vector and score is often highly non-linear 262 and generator-dependent. For instance, the generated images often vary across the prompts and 263 different T2I models, which can have substantial effect on the resulting CLIPScore (1). Second, the 264 kernel methods can approximate any function arbitrarily well with enough training data and enjoy 265 nice statistical properties.

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267 5.1 THE SCK-UCB ALGORITHM

In this section, we present SCK-UCB in Algorithm 2, an optimism-based approach to the online 269 evaluation of prompt-base generation. At each iteration, SCK-UCB first estimates the scores via kernel ridge regression (KRR) and then picks the model with the highest estimated score. To construct the KRR dataset, the algorithm maintains index sets $\{\Psi_g\}_{g \in \mathcal{G}}$, where each set $\Phi_g \subseteq [T]$ stores the iterations such that model g is chosen (line 6).

273 274 Algorithm 2 Shared-Context Kernel UCB (SCK-UCB) 275 **Require:** total iterations $T \in \mathbb{N}_+$, set of generators $\mathcal{G} = [G]$, prompt distribution $\rho \in \Delta(\mathcal{Y})$, score 276 function $s: \mathcal{Y} \times \mathcal{X} \to [-1, 1]$, positive definite kernel $k: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$, regularization and 277 exploration parameters $\alpha, \eta \geq 0$ 278 **Initialize:** observation sequence $\mathcal{D} \leftarrow \emptyset$ and index set $\Psi_q \leftarrow \emptyset$ for all $g \in \mathcal{G}$ 279 1: for iteration $t = 1, 2, \cdots, T$ do 2: Prompt $y_t \sim \rho$ is revealed. Compute $\{(\widehat{\mu}_g, \widehat{\sigma}_g) \leftarrow \text{COMPUTE}_UCB(\mathcal{D}, y_t, \Psi_g)\}_{g \in \mathcal{G}}$. 3: 281 4: Pick model $g_t \leftarrow \arg \max_{g \in \mathcal{G}} \{ \widehat{s}_g \}$, where $\widehat{s}_g \leftarrow \widehat{\mu}_g + (2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_g$. 282 Sample an answer $x_t \sim P_{g_t}(\cdot|y_t)$ and compute the score $s_t \leftarrow s(y_t, x_t)$. Update $\mathcal{D} \leftarrow \mathcal{D} \cup \{(y_t, s_t)\}$ and $\Psi_{g_t} \leftarrow \Psi_{g_t} \cup \{t\}$. 5: 283 6: 284 7: end for 285 286 8: function COMPUTE_UCB(\mathcal{D}, y, Ψ_a) 287 $\begin{array}{l} \text{if } \Psi_g \text{ is empty then} \\ \widehat{\mu}_g \leftarrow +\infty, \widehat{\sigma}_g \leftarrow +\infty. \end{array}$ 9: 288 10: 289 11: else Set $K \leftarrow [k(y_i, y_j)]_{i,j \in \Psi_g}, v \leftarrow [s_i]_{i \in \Psi_g}^{\top}$, and $k_y \leftarrow [k(y, y_i)]_{i \in \Psi_g}^{\top}$. 290 12: 291 $\widehat{\mu}_g \leftarrow k_y^\top (K + \alpha I)^{-1} v.$ 13: 292 $\widehat{\sigma}_g \leftarrow \alpha^{-\frac{1}{2}} \sqrt{k(y,y) - k_y^\top (K + \alpha I)^{-1} k_y}.$ 14: 293 15: end if 16: return $(\widehat{\mu}_g, \widehat{\sigma}_g)$. 295 17: end function 296

The key design in SCK-UCB is the function COMPUTE_UCB (lines 8-17), which outputs both the KRR estimator $\hat{\mu}_g$ and an uncertainty quantifier $\hat{\sigma}_g$. The estimated score is then computed by $\hat{s}_g = \hat{\mu}_g + (2\eta + \sqrt{\alpha})\hat{\sigma}_g$ (line 4), which is initially set to $+\infty$ to ensure each model is picked at least once (lines 9-10). Particularly, the following lemma shows that under some conditions, \hat{s}_g is an optimistic estimation of $s_g(y_t)$ with high probability. The detailed proof can be found in Appendix C.1.

Lemma 1 (Optimism). Let $\Psi_g \subseteq [T]$ be an index set such that the set of scores $\{s_t : t \in \Psi_g\}$ are independent random variables. Then, under Assumption 1, with probability at least $1 - \delta$, the quantity $\hat{\mu}_g$ computed in function COMPUTE_UCB(\mathcal{D}, y, Ψ_g) satisfies that

$$|\widehat{\mu}_g - s_g(y)| \le (2\eta + \sqrt{\alpha})\widehat{\sigma}_g,\tag{5}$$

where $\eta = \sqrt{2\log(2/\delta)}$. Hence, it holds that $\hat{s}_g = \hat{\mu}_g + (2\eta + \sqrt{\alpha})\hat{\sigma}_g \ge s_g(y)$.

We show that a variant of SCK-UCB attains a regret of $O(\sqrt{GT})$. The formal statement and the proof can be found in Appendix A.

Theorem 1 (Regret, informal). Under the same conditions in Lemma 1, with probability of at least $1 - \delta$, a variant of Algorithm 2 attains a regret of $\tilde{O}(\sqrt{GT})$.

3165.2SCK-UCB with RANDOM FOURIER FEATURES

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The SCK-UCB solves a KRR for each model at an iteration to estimate the scores, which can be expensive in both computation and memory for a large number of iterations. To address this problem, we leverage the random Fourier features (RFF) sampling (Rahimi & Recht, 2007a) for positive definite shift-invariant kernels. At a high level, RFF maps the input data, e.g., the prompt (vector) in our setting, to a randomized low-dimensional feature space and then applies fast linear methods to solve the regression problem. Particularly, the inner product between these projected randomized features is an *unbiased* estimation of the kernel value.

324 We present the RFF-UCB algorithm, which is a variant of SCK-UCB with random features. RFF-325 UCB leverages a RFF-based approach to compute the mean and uncertainty quantifier in line 3 of 326 Algorithm 2, which we present in Algorithm 3. Upon receiving the regression dataset consisting of 327 prompt-score pairs, COMPUTE_UCB_RFF first projects each d-dimensional prompt vector to a ran-328 domized s-dimensional feature space according to Equation (3) (lines 5-6) and then solves a linear ridge regression to estimate the mean and uncertainty (lines 8-9). To derive statistical guarantees, the number of features varies according to both the input data and error thresholds, which we will 330 specify in Appendix B.1. In practice, we find that a size around 50 can attain satisfactory empirical 331 performance. To see why RFF can reduce the computation, note that the size of the (regularized) 332 Gram matrix $(\Phi_g^{\top} \Phi_g + \alpha I)$ in line 8 is fixed to be s in the whole process, while the size of $(K + \alpha I)$ 333 in line 13 of Algorithm 2 scales with $|\Psi_q|$ and can grow linearly over iterations. Particularly, the 334 following lemma shows that COMPUTE_UCB_RFF can reduce the time and space by an order of 335 $O(t^2)$ and O(t), respectively. 336

Algorithm 3 Compute UCB with Random Fourier Features

Require: the Fourier transform p of a positive definite shift-invariant kernel k(y, y') = k(y - y'), 339 error thresholds ϵ_{RFF} , $\Delta_{\text{RFF}} > 0$, regularization and exploration parameters $\alpha, \eta \ge 0$ 340 **Initialize:** number of features s, bonus terms $\mathcal{B}_{q,1}$ and $\mathcal{B}_{q,2}$ 341 1: **function** COMPUTE_UCB_RFF(\mathcal{D}, y, Ψ_g) 342 $\begin{array}{l} \text{if } \Psi_g \text{ is empty then} \\ \widetilde{\mu}_g \leftarrow +\infty, \widetilde{\sigma}_g \leftarrow +\infty. \\ \text{else} \end{array}$ 2: 343 3: 344 4: Draw $\omega_1, \cdots, \omega_s \stackrel{\text{i.i.d.}}{\sim} p$ and $b_1, \cdots, b_s \stackrel{\text{i.i.d.}}{\sim} \text{Unif}([0, 2\pi]).$ 345 5: Define mapping $\varphi(y') \leftarrow \sqrt{\frac{2}{s}} \cdot [\cos(w_1^\top y + b_1), \cdots, \cos(w_s^\top y + b_s)]^\top$ for any $y' \in \mathbb{R}^d$. 6: 347 Set $\widetilde{\Phi}_g \leftarrow [\varphi(y_i)^\top]_{i \in \Psi_g}$ and $v \leftarrow [s_i]_{i \in \Psi_g}^\top$. 7: 348
$$\begin{split} &\widetilde{\mu}_g \leftarrow (\varphi(y))^\top (\widetilde{\Phi}_g^\top \widetilde{\Phi}_g + \alpha I)^{-1} \widetilde{\Phi}_g^\top v + \mathcal{B}_{g,1}. \\ &\widetilde{\sigma}_g \leftarrow \alpha^{-\frac{1}{2}} \sqrt{1 - (\varphi(y))^\top (\widetilde{\Phi}_g^\top \widetilde{\Phi}_g + \alpha I)^{-1} \widetilde{\Phi}_g^\top \widetilde{\Phi}_g(\varphi(y))} + \mathcal{B}_{g,2}. \end{split}$$
349 8: 350 9: 351 10: end if 352 return $(\widetilde{\mu}_g, \widetilde{\sigma}_g)$. 11: 353 12: end function 354 355

Lemma 2 (Time and space complexity). At any iteration $t \in [T]$, COMPUTE_UCB (lines 8-17 of Algorithm 2) requires $O(t^3/G^2)$ time and $O(t^2/G)$ space, while COMPUTE_UCB_RFF with random features of size $s \in \mathbb{N}_+$ (Algorithm 3) requires $O(ts^2)$ time and O(ts) space, where G is the number of generators. See Appendix B.5 for details.

It can be shown that the implementation of SCK-UCB with RFF attains the exact same regret guarantees for adaptively selected feature sizes. The formal statement and the proof can be found in Appendix B.2.

Theorem 2 (Regret when using RFF, informal). Under the same conditions in Theorem 4, a variant of RFF-UCB attains a regret of $\widetilde{O}(\sqrt{GT})$.

366 6 NUMERICAL RESULTS

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In this section, we provide numerical results for the proposed SCK-UCB-poly3 algorithm (SCK-UCB using polynomial kernel with degree 3, i.e., $k_{\text{poly3}}(x_1, x_2) = (1 + x_1^{\top} x_2)^3)$ and the RFF-UCB algorithm using RBF kernel on various prompt-based generation tasks, including text-to-image (T2I) generation, image captioning (image-to-text), and text-to-video (T2V) generation.

[Baselines. We compare the proposed methods with five baselines, including 1) Lin-UCB: SCK-UCB with linear kernel, i.e., $k_{lin}(x_1, x_2) = x_1^{\top} x_2$, which does not incorporate non-linearity in score estimation, 2) One-arm Oracle: always picking the model with the maximum averaged CLIPScore, 3) Naive-KRR: SCK-UCB-poly3 without exploration, which selects the model with the highest estimated mean conditioned to the prompt, 4) Greedy: always generating samples from the model with the highest empirical CLIPScore, and 5) Random: uniformly selecting a generator at each step. The results for baselines are presented by dot lines with different colors.] Performance metrics. For each experiment, we report three performance metrics: [(i) *outscore-the-best* (O2B): the difference between the CLIPScore attained by the algorithm and the highest average CLIPScore attained by any single model], (ii) *optimal-pick-ratio* (OPR): the overall ratio that the algorithm picks the best generator conditioned to the prompt type, (iii) *moving-average OPR*: OPR over the last 100 iterations.

383 [Summary of results. The main finding of our numerical experiments is the improvement of the 384 proposed contextual bandit SCK-UCB algorithm over the one-arm oracle baseline. This result means 385 that the online learning algorithm can outperform a user with side-knowledge of the single best-386 performing model, which is made possible by a prompt-based selection of the model. This finding supports the application of contextual bandit algorithms in the selection of text-based generative 387 388 models. Moreover, our numerical results indicate that the proposed SCK-UCB algorithm can perform better with a non-linear kernel function. Finally, in our experiments, the proposed RFF-UCB 389 variant could reduce the computational costs of the general SCK-UCB algorithm.] 390

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6.1 TEXT-TO-IMAGE GENERATION

Setup 1: Prompt-based selection between real generative models. The first set of experiments are on the setup illustrated in Figure 1, where we generate images from two T2I generators, including Stable Diffusion v1-5¹ and PixArt- α -XL-2-512x512² (see Figure 2). The results show that SCK-UCB-poly3 outperforms the baseline algorithm and attains a high optimal-pick-ratio, which shows that it can identify the optimal model conditioned to the prompt. Additionally, we provide numerical results on various T2I generative models, including uni-Diffuser³ and DeepFloyd IF-I-M-v1.0.⁴ (see Figures 14, 15, and 16 in the Appendix).



Figure 2: Prompt-based selection between Stable Diffusion and PixArt- α (Setup 1): Results are averaged over 20 trials.

[Setup 2: Adapt to newly-introduced prompts and generators. We consider scenarios where new generative models or prompt types are introduced after the initial deployment. In the first experiment, there are two available generators initially, including Stable Diffusion and PixArt- α . After 2,500 iterations, uniDiffuser is also available (see Figure 3). In the second experiment, we generate samples from both PixArt- α and uniDiffuser, and a new prompt type is introduced after each 1,000 iterations (see Figure 13 in the Appendix). The results show that SCK-UCB-poly3 can well adapt to new prompt types and generators.]

Setup 3: Synthetic expert T2I models. In this setup, we synthesize five T2I generators based on Stable Diffusion 2, where each generator is an "expert" in generating images corresponding to a prompt type. The prompts are captions in the MS-COCO dataset from five categories: dog, car, carrot, cake, and bowl. At each iteration, a caption is drawn from a (random) category, and an image

^{429 &}lt;sup>1</sup>https://huggingface.co/docs/diffusers/en/api/pipelines/stable_diffusion/text2img

^{430 &}lt;sup>2</sup>https://huggingface.co/PixArt-alpha/PixArt-XL-2-512x512

^{431 &}lt;sup>3</sup>https://github.com/thu-ml/unidiffuser

⁴https://github.com/deep-floyd/IF



Figure 3: Adapt to newly-introduced generators (Setup 2): Results are averaged over 20 trials.

is generated from Stable Diffusion 2. If the learner does not select the expert generator, then we add Gaussian noise to the generated image. Examples are visualized in Figure 4.



467 Figure 4: Generated images with noise perturbations: Each row and column display the generated
468 images from a synthetic generator according to one single type of prompts. Images generated by the
469 expert models are framed by green boxes. Gaussian noises are applied to non-expert models.



Figure 5: Synthetic expert T2I models (Setup 3): Results are averaged over 20 trials.

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486 6.2 **RESULTS ON OTHER PROMPT-BASED GENERATION TASKS** 487

488 Setup 4: Image Captioning. In this setup, the images are chosen from the MS-COCO dataset 489 from five categories: dog, car, carrot, cake, and bowl. Similar to Section 6.1, we synthesize five expert generators based on vit-gpt2 model in the Transformers repository.⁵ If a non-expert generator is 490 chosen, then the caption is generated from the noisy image perturbed by Gaussian noises. Examples are visualized in Figure 19. The numerical results are summarized in Figure 6.



Figure 6: Image captioning (Setup 4): Results are averaged over 20 trials.

Setup 5: Synthetic Text-to-Video (T2V) task. We provide numerical results on a synthetic T2V setting. Specifically, both the captions and videos are randomly selected from the following five categories of the MSR-VTT dataset (Xu et al., 2016): sports/action, movie/comedy, vehicles/autos, music, and food/drink. Each of the five synthetic arms corresponds to an expert in "generating" videos from a single category. Gaussian noises are applied to the video for non-experts. The results are summarized in Figure 7.



Figure 7: Synthetic T2V task (Setup 5): Results are averaged over 20 trials.

7 CONCLUSION

533 In this work, we investigated prompt-based selection of generative models using a contextual ban-534 dit algorithm, which can identify the best available generative model for a given text prompt. 535 We adapted the Kernel-UCB algorithm to perform this selection task and proposed two new al-536 gorithms: SCK-UCB and RFF-UCB. Our numerical results on text-to-image, text-to-video, and 537 image-captioning tasks demonstrate the effectiveness of the proposed framework in scenarios where 538 the available generative models have varying performance rankings depending on the type of prompt.

⁵https://huggingface.co/nlpconnect/vit-gpt2-image-captioning

An interesting direction for future research is to extend the application of our algorithms to text-to-text language models, where different models may respond better to questions on different topics.
 Furthermore, considering evaluation criteria beyond relevance, such as diversity and novelty scores, could lead to extensions of our proposed framework.

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756 A PROOF IN SECTION 5.1

The technical challenge in analyzing SCK-UCB is that predictions in later iterations are made use of previous outcomes. Hence, the rewards $\{s_i\}_{i \in \Psi_g}$ are not independent if the index set Φ_g is updated each time when model g is chosen (line 6 of Algorithm 2). To address this problem, we leverage a standard approach used in prior works (Auer, 2003; Chu et al., 2011; Valko et al., 2013) and present a variant of SCK-UCB in Algorithm 4, which is called Sup-SCK-UCB. We prove it attains a regret of order $\widetilde{O}(\sqrt{T})$.

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766 767 A.1 THE Sup-SCK-UCB ALGORITHM

Algorithm 4 Sup-SCK-UCB

768	Rec	puire: total iterations $T \in \mathbb{N}_+$, set of generators $\mathcal{G} = [G]$, prompt distribution $\rho \in \Delta(\mathcal{Y})$, score
769		function $s: \mathcal{Y} \times \mathcal{X} \to [-1, 1]$, positive definite kernel $k: \mathcal{Y} \times \mathcal{Y} \to \mathbb{R}$, regularization and
770		exploration parameters $\alpha, \eta \ge 0$, function COMPUTE_UCB in Algorithm 2
771	Init	ialize: observation sequence $\mathcal{D} \leftarrow \emptyset$ and index sets $\{\Psi_q^m \leftarrow \emptyset\}_{m=1}^M$ for all $g \in \mathcal{G}$, where
772		$M \leftarrow \log T$
773	1:	for iteration $t = 1, 2, \cdots, T$ do
774	2:	Prompt $y_t \sim \rho$ is revealed.
775	3:	Set stage $m \leftarrow 1$ and $\mathcal{G}^1 \leftarrow \mathcal{G}$.
776	4:	repeat
777	5:	Compute $\{(\widehat{\mu}_{g}^{m}, \widehat{\sigma}_{g}^{m}) \leftarrow \text{COMPUTE}_{-}\text{UCB}(\mathcal{D}, y_{t}, \Psi_{g}^{m})\}_{g \in \widehat{\mathcal{G}}^{m}}.$
778	6:	Set upper confidence bound $\widehat{s}_g^m(y) \leftarrow \widehat{\mu}_g^m + (2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_g^m$ for all $g \in \widehat{\mathcal{G}}^m$.
779	7:	if $(2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_a^m \leq 1/\sqrt{T}$ for all $g \in \widehat{\mathcal{G}}^m$ then
780	8:	Pick model $g_t \leftarrow \arg \max \widehat{s}_g^m(y)$.
782	9:	else if $(2\eta+\sqrt{lpha})\cdot\widehat{\sigma}_g^m\leq 2^{-m}$ for all $g\in\widehat{\mathcal{G}}^m$ then
783	10:	$\widehat{\mathcal{G}}^{m+1} \leftarrow \{g \in \widehat{\mathcal{G}}^m : \widehat{s}_g^m(y) \ge \max_{g \in \widehat{\mathcal{G}}^m} \widehat{s}_g^m(y) - 2^{1-m}\}.$
784	11:	Set stage $m \leftarrow m + 1$.
785	12:	else
786	13:	Pick $g_t \in \mathcal{G}^m$ such that $(2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_g^m > 2^{-m}$.
787	14:	Update $\Psi_{g_t}^m \leftarrow \Psi_{g_t}^m \cup \{t\}.$
788	15:	end if
789	16:	until a model g_t is selected
790	17:	Sample an answer $x_t \sim P_{g_t}(\cdot y_t)$ and compute the score $s_t \leftarrow s(y_t, x_t)$.
791	18:	Update $\mathcal{D} \leftarrow \mathcal{D} \cup \{(y_t, s_t)\}.$
700	19:	end for

A.2 ANALYSIS

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In this section, we prove the following regret bound of Sup-SCK-UCB. **Theorem 3** (Regret of Sup-SCK-UCB). Under Assumption 1, with probability at least $1 - \delta$, the regret of Sup-SCK-UCB with $\eta = \sqrt{2 \log(2(\log T)GT/\delta)}$ is bounded by

$$\operatorname{Regret}(T) \le \widetilde{O}\left((1+\sqrt{\alpha})\sqrt{d_{\operatorname{eff}}\left(10+\frac{15}{\alpha}\right)GT}\right)$$
(6)

where d_{eff} is a data-dependent quantity defined in Lemma 10 and logarithmic factors are hidden in the notation $\widetilde{O}(\cdot)$.

Notations. To facilitate the analysis, we add an subscript t to all the notations in Algorithm 4 to indicate they are quantities computed at the t-th iteration, i.e., $\hat{\mu}_{g,t}^m$, $\hat{\sigma}_{g,t}^m$, $\hat{s}_{g,t}^m$, \hat{g}_t^m , and $\Psi_{g,t}^m$.⁶

⁶Note that line 14 in the algorithm is rewritten as " $\Psi_{g_t,t+1}^m \leftarrow \Psi_{g_t,t+1}^m \cup \{t\}, \Psi_{g_t,t+1}^{m'} \leftarrow \Psi_{g_t,t}^{m'}$ for any $m' \neq m$, and $\Psi_{g,t+1}^{m'} \leftarrow \Psi_{g,t}^{m'}$ for any $g \neq g_t$ and $m \in [M]$ ". In addition, we set $\Psi_{g,t+1}^m \leftarrow \Psi_{g,t+1}^m$ for all $g \in \mathcal{G}$ and $m \in [M]$ in line 8.

We leverage the following two lemmas to prove Theorem 3. The first lemma shows that the construction of index sets $\{\Psi_{g,t}^m\}_{m=1}^M$ ensures the independence among the rewards $\{s_i\}_{i \in \Psi_{g,t}^m}$, which allows us to utilize Lemma 1 to bound the estimation error.

Lemma 3 (Auer (2003), Lemma 14). For any iteration $t \in [T]$, model $g \in \mathcal{G}$, and stage $m \in [M]$, the set of rewards $\{s_i\}_{i \in \Psi_{g,t}^m}$ are independent random variables such that $\mathbb{E}[s_i] = s_g(y_i)$.

The second lemma shows several properties of the estimated score $\hat{s}_{g,t}^m$ and the set $\hat{\mathcal{G}}_t^m$. The detailed proof can be found in Appendix C.2.

Lemma 4 (Valko et al. (2013), Lemma 7). With probability at least $1 - (MGT)\delta$, for any iteration $t \in [T]$ and stage $m \in [M]$, the following hold:

•
$$|\widehat{\mu}_{a,t}^m - s_q(y_t)| \leq (2\eta + \sqrt{\alpha})\widehat{\sigma}_{a,t}^m$$
 for any $g \in \widehat{\mathcal{G}}_t^m$,

•
$$\arg \max_{a \in G} s_a(y_t) \in \mathcal{G}_t^m$$
, and

•
$$s_{\star}(y_t) - s_g(y_t) \le 2^{3-m}$$
 for any $g \in \widehat{\mathcal{G}}_t^m$.

Now, we are ready to finish the proof of Theorem 3.

Proof of Theorem 3. Let $\mathcal{T}_1 := \bigcup_{m \in [M], g \in \mathcal{G}} \Psi_{g,T+1}^m$ and $\mathcal{T}_0 := [T] \setminus \mathcal{T}_1$. Note that \mathcal{T}_0 and \mathcal{T}_1 are sets of iterations such that model is picked in lines 8 and 13 of Algorithm 4, respectively.

1. Regret incurred in \mathcal{T}_0 . For any $t \in [T]$, let m_t denote the stage that model g_t is picked at the *t*-th iteration. We have that

$$\sum_{t\in\mathcal{T}_{0}} (s_{\star}(y_{t}) - s_{g_{t}}(y_{t})) \leq \sum_{t\in\mathcal{T}_{0}} (\widehat{s}_{g_{\star,t},t}^{m_{t}}(y) - s_{g_{t}}(y_{t}))$$

$$\leq \sum_{t\in\mathcal{T}_{0}} (\widehat{s}_{g_{t},t}^{m_{t}}(y) - s_{g_{t}}(y_{t}))$$

$$= \sum_{t\in\mathcal{T}_{0}} \left(\widehat{\mu}_{g_{t},t}^{m_{t}} + (2\eta + \sqrt{\alpha})\widehat{\sigma}_{g_{t},t}^{m_{t}} - s_{g_{t}}(y_{t})\right)$$

$$\leq 2(2\eta + \sqrt{\alpha}) \sum_{t\in\mathcal{T}_{0}} \widehat{\sigma}_{g_{t},t}^{m_{t}}$$

$$\leq 2(2\eta + \sqrt{\alpha}) \sum_{t\in\mathcal{T}_{0}} T^{-\frac{1}{2}} \leq 2(2\eta + \sqrt{\alpha})\sqrt{T},$$
(7)

where the second inequality holds by the definition of g_t and the fact that $g_{\star,t} \in \widehat{\mathcal{G}}_t^{m_t}$, and the fifth inequality holds by line 7 of Algorithm 4.

2. Regret incurred in T_1 .

$$\sum_{t \in \mathcal{T}_{1}} (s_{\star}(y_{t}) - s_{g_{t}}(y_{t})) = \sum_{g \in \mathcal{G}} \sum_{m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} (s_{\star}(y_{t}) - s_{g_{t}}(y_{t}))$$

$$\leq \sum_{g \in \mathcal{G}} \sum_{m \in [M]} 2^{3-m} \cdot |\Psi_{g,T+1}^{m}|, \qquad (8)$$

where the inequality holds by the last statement in Lemma 4. It remains to bound $|\Psi_{g,T+1}^{m}|$. First note that for any $m \in [M]$, we have that

$$(2\eta + \sqrt{\alpha}) \sum_{t \in \Psi^m_{g,T+1}} \widehat{\sigma}^m_{g,t} > 2^{-m} \cdot |\Psi^m_{g,T+1}|$$

from line 13 of Algorithm 4. In addition, by a similar statement of (Valko et al., 2013, Lemma 4), which is stated in Lemma 10, we have that

$$\sum_{t \in \Psi_{g,T+1}^m} \widehat{\sigma}_{g,t}^m \le \widetilde{O}\left(\sqrt{d_{\text{eff}}\left(10 + \frac{15}{\alpha}\right)|\Psi_{g,T+1}^m|}\right),\tag{9}$$

where d_{eff} is defined therein and logarithmic factors are hidden in the notation $O(\cdot)$. Plugging in Equation (8) results in

$$\sum_{t \in \mathcal{T}_{1}} (s_{\star}(y_{t}) - s_{g_{t}}(y_{t})) \leq \widetilde{O}\left((1 + \sqrt{\alpha}) \sum_{g \in \mathcal{G}} \sum_{m \in [M]} \sqrt{d_{\text{eff}}\left(10 + \frac{15}{\alpha}\right) |\Psi_{g,T+1}^{m}|}\right)$$
$$\leq \widetilde{O}\left((1 + \sqrt{\alpha})\sqrt{GM} \sqrt{d_{\text{eff}}\left(10 + \frac{15}{\alpha}\right) \sum_{g \in \mathcal{G}} \sum_{m \in [M]} |\Psi_{g,T+1}^{m}|}\right) \quad (10)$$
$$\leq \widetilde{O}\left((1 + \sqrt{\alpha}) \sqrt{d_{\text{eff}}\left(10 + \frac{15}{\alpha}\right) GT}\right),$$

where the second inequality holds by Cauchy-Schwarz inequality.

3. Putting everything together. Combining Inequalities (7) and (10) leads to

$$\operatorname{Regret}(T) = \left(\sum_{t \in \mathcal{T}_0} + \sum_{t \in \mathcal{T}_1}\right) \left(s_\star(y_t) - s_{g_t}(y_t)\right) \le \widetilde{O}\left(\left(1 + \sqrt{\alpha}\right)\sqrt{d_{\operatorname{eff}}\left(10 + \frac{15}{\alpha}\right)GT}\right),$$

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which concludes the proof.

PROOF IN SECTION 5.2 В

B.1 ERROR OF KRR ESTIMATORS WITH RANDOM FEATURES

Theorem 4. Assume the error thresholds input to Algorithm 3 satisfy that Δ_{RFF} , $\epsilon_{\text{RFF}} \leq 1/2$. Under the same conditions in Lemma 1, with probability at least $1-2\delta$, the quantity $\tilde{\mu}_q$ computed by function COMPUTE_UCB_RFF satisfies that

$$|\widetilde{\mu}_g - s_g(y)| \le \mathcal{B}_{g,1} + (2\eta + \sqrt{\alpha})(\widetilde{\sigma}_g + \mathcal{B}_{g,2}),$$

with the number of features satisfying Inequality (12) and bonus terms $\mathcal{B}_{g,1}$ and $\mathcal{B}_{g,2}$ given by Equations (11) and (13), where $\eta = \sqrt{2\log(2/\delta)}$. Hence, it holds that $\tilde{s}_g = \tilde{\mu}_g + \mathcal{B}_{g,1} + (2\eta + 2\eta)$ $\sqrt{\alpha}$) $(\widetilde{\sigma}_g + \mathcal{B}_{g,2}) \ge s_g(y).$

Proof. The proof is based on the following two lemmas, which analyze the concentration error of the quantities $\tilde{\mu}_g$ and $\tilde{\sigma}_g$. The detailed proof can be found in Appendix B.3 and B.4, respectively.

Lemma 5 (Concentration of mean using RFF). Let Δ_{RFF} , $\epsilon_{\text{RFF}} \leq 1/2$. Under the same condi-tions in Lemma 1, with probability at least $1 - \delta$, the quantity $\tilde{\mu}_g$ computed by function COM-PUTE_UCB_RFF satisfies that

$$|\widetilde{\mu}_g - \widehat{\mu}_g| \le \mathcal{B}_{g,1} := \alpha^{-1} |\Psi_g| \,\epsilon_{\text{RFF}} + \alpha^{-2} \,\Delta_{\text{RFF}}(||K||_2 + \alpha) \tag{11}$$

with number of features

$$s \ge \max\left\{\frac{4(d+2)}{\epsilon_{\mathsf{RFF}}^2} \log\left(\frac{\sigma_p^2}{(\delta/2) \cdot \epsilon_{\mathsf{RFF}}^2} \cdot 2^8\right), \frac{8|\Psi_g|}{3\alpha} \Delta_{\mathsf{RFF}}^{-2} \log\left(\frac{32s_\alpha(K)}{\delta}\right)\right\},\tag{12}$$

where $\hat{\mu}_g = (\phi(y))^\top \Phi_g^\top (K + \alpha I)^{-1} v$, and σ_p^2 and $s_\alpha(\cdot)$ are two quantities defined in Lemmas 8 and 9, respectively.

Lemma 6 (Concentration of variance using RFF). Let c > 0 denote a lower bound of 1 - $\|k_y\|_{(K+\alpha I)^{-1}}^2$. Then, conditioned on the successful events in Lemma 5, the quantity $\widetilde{\sigma}_g$ computed by function COMPUTE_UCB_RFF satisfies that

$$|\widetilde{\sigma}_g - \widehat{\sigma}_g| \le \mathcal{B}_{g,2} := (c \cdot \alpha)^{-\frac{1}{2}} \left(2|\Psi_g| \alpha^{-2} \Delta_{\mathsf{RFF}}(\|K\|_2 + \alpha) + 3\alpha^{-1} |\Psi_g| \epsilon_{\mathsf{RFF}} \right)$$
(13)

where
$$\widehat{\sigma}_g = \alpha^{-\frac{1}{2}} \sqrt{k(y,y) - k_y^\top (K + \alpha I)^{-1} k_y}$$
.

Finally, combining Lemmas 1, 5, and 6, we derive that

$$\begin{aligned} |\widetilde{\mu}_{g} - s_{g}(y)| \leq & |\widetilde{\mu}_{g} - \widehat{\mu}_{g}| + |\widehat{\mu}_{g} - s_{g}(y)| \\ \leq & \mathcal{B}_{g,1} + (2\eta + \sqrt{\alpha})\widehat{\sigma}_{g} \\ \leq & \mathcal{B}_{g,1} + (2\eta + \sqrt{\alpha})(\widetilde{\sigma}_{g} + \mathcal{B}_{g,2}), \end{aligned}$$

which concludes the proof.

B.2 Sup-SCK-UCB WITH RANDOM FOURIER FEATURES

Algorithm description. To apply RFF to Sup-SCK-UCB, we replace function COMPUTE_UCB with COMPUTE_UCB_RFF in Algorithm 4. To achieve the regret bound (6), an important problem is to design (adaptive) error thresholds, i.e., ϵ_{RFF} and Δ_{RFF} , when computing UCB at each stage m and iteration t. We prove the regret bound in the following theorem.

Theorem 5 (Regret of Sup-RFF-UCB). Under Assumption 1, with probability at least $1 - \delta$, Sup-RFF-UCB attains the regret bound (6), where $\eta = \sqrt{2\log(4(\log T)GT/\delta)}$ and sequence of error thresholds input to function COMPUTE_UCB_RFF satisfying Equation (14).

Proof. The proof is similar to the proof of Theorem 3. First, combining Lemmas 8 and 3, the following lemma can be proved by the exact same analysis for Lemma 4.

Lemma 7. With probability at least $1 - (2MGT)\delta$, for any iteration $t \in [T]$ and stage $m \in [M]$, the following hold:

•
$$|\widetilde{\mu}_{a,t}^m - s_q(y_t)| \leq \mathcal{B}_{a,1,t}^m + (2\eta + \sqrt{\alpha})(\widetilde{\sigma}_{a,t}^m + \mathcal{B}_{a,2,t}^m)$$
 for any $g \in \widehat{\mathcal{G}}_t^m$,

• $\operatorname{arg\,max}_{a \in \mathcal{G}} s_q(y_t) \in \widehat{\mathcal{G}}_t^m$, and

 $\sum \left(s_{\star}(y_t) - s_{g_t}(y_t) \right)$

•
$$s_{\star}(y_t) - s_q(y_t) \leq 2^{3-m}$$
 for any $g \in \widehat{\mathcal{G}}_t^m$

where the first statement is guaranteed by Theorem 4, $\mathcal{B}_{g,1,t}^m$ and $\mathcal{B}_{g,2,t}^m$ are the bonus (11) and (13) computed at the *m*-th stage of iteration *t*.

Next, for iterations in \mathcal{T}_0 (model g_t is picked in line 8 of Algorithm 4), we still have

$$\sum_{t \in \mathcal{T}_0} (s_\star(y_t) - s_{g_t}(y_t)) \le \widetilde{O}(\sqrt{\alpha T})$$

Further, for iterations in T_1 (model g_t is picked in line 13 of Algorithm 4), the third statement in the above lemma and line 14 of Algorithm 3 ensure that

$$\begin{split} & = \sum_{q \in \mathcal{G}} \sum_{m \in [M]} 2^{3-m} \cdot |\Psi_{g,T+1}^{m}| \\ & \leq 8 \sum_{g \in \mathcal{G}} \sum_{m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} \left(\mathcal{B}_{g_{t},1,t}^{m_{t}} + (2\eta + \sqrt{\alpha}) \left(\widetilde{\sigma}_{g,t}^{m} + \mathcal{B}_{g_{t},2,t}^{m_{t}} \right) \right) \\ & \leq 8 \sum_{g \in \mathcal{G}} \sum_{m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} \left(\mathcal{B}_{g_{t},1,t}^{m_{t}} + (2\eta + \sqrt{\alpha}) \left(\widehat{\sigma}_{g,t}^{m} + 2\mathcal{B}_{g_{t},2,t}^{m_{t}} \right) \right) \\ & = 8 \sum_{g \in \mathcal{G}} \sum_{m \in [M]} \left(\sum_{t \in \Psi_{g,T+1}^{m}} \left(\mathcal{B}_{g_{t},1,t}^{m_{t}} + 2(2\eta + \sqrt{\alpha}) \mathcal{B}_{g_{t},2,t}^{m_{t}} \right) + (2\eta + \sqrt{\alpha}) \sum_{t \in \Psi_{g,T+1}^{m}} \widehat{\sigma}_{g,t}^{m} \right). \end{split}$$

970971 Note that the upper bound of the second term has been derived in Equation (9). It remains to bound the first term. Essentially, we will find a sequence of error thresholds, and hence the number of

features defined in Inequality (12), such that the first term is bounded by $\widetilde{O}(\sqrt{GT})$. For convenience, we introduce the following notations:

Additional notations. For any iteration $t \in [T]$, model $g \in \mathcal{G}$, and stage $m \in [M]$, we define $K_{g,t}^m := \Phi_{g,t}^m (\Phi_{g,t}^m)^\top$, where $\Phi_{g,t}^m := [\phi(y_i)^\top]_{i \in \Psi_{g,t}^m}$. In addition, we denote by $0 < c_{g,t}^m \le 1$ the lower bound in Lemma 6 corresponding to y_t and $K_{g,t}^m$. Let

$$\epsilon_{\mathsf{RFF},g,t}^{m} \le t^{-\frac{1}{2}} (|\Psi_{g,t}^{m}|)^{-1} \sqrt{G \cdot c_{g,t}^{m}}, \quad \Delta_{\mathsf{RFF},g,t}^{m} \le t^{-\frac{1}{2}} (|\Psi_{g,t}^{m}| (\|K_{g,t}^{m}\|_{2} + \alpha))^{-1} \sqrt{G \cdot c_{g,t}^{m}}$$
(14)

denote the (upper bound of) error thresholds input to function COMPUTE_UCB_RFF.

$$\sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} \mathcal{B}_{g_{t},1,t}^{m}$$

$$= \sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} \left(\alpha^{-1} | \Psi_{g,t}^{m} | \epsilon_{\mathsf{RFF},g,t}^{m} + \alpha^{-2} \Delta_{\mathsf{RFF},g,t}^{m} (\| K_{g,t}^{m} \|_{2} + \alpha) \right)$$

$$\leq \sqrt{G} \sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g,T+1}^{m}} \left(\alpha^{-1} t^{-\frac{1}{2}} + \alpha^{-2} t^{-\frac{1}{2}} \right)$$

$$\leq \sqrt{G} \sum_{t=1}^{T} \left(\alpha^{-1} t^{-\frac{1}{2}} + \alpha^{-2} t^{-\frac{1}{2}} \right)$$

$$\leq O\left(\left(\alpha^{-1} + \alpha^{-2} \right) \sqrt{GT} \right)$$
(15)

where the first inequality holds by the fact that each $t \in [T]$ appears in at most one index set.

$$\begin{split} &\sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g, T+1}^{m}} \mathcal{B}_{g_{t}, 2, t}^{m} \\ &\leq \sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g, T+1}^{m}} (c \cdot \alpha)^{-\frac{1}{2}} \left(2|\Psi_{g}| \alpha^{-2} \Delta_{\mathsf{RFF}, g, t}^{m} (\|K_{g, t}^{m}\|_{2} + \alpha) + 3\alpha^{-1} |\Psi_{g, t}^{m}| \epsilon_{\mathsf{RFF}, g, t}^{m} \right) \\ &\leq \sqrt{G} \sum_{g \in \mathcal{G}, m \in [M]} \sum_{t \in \Psi_{g, T+1}^{m}} \left(2\alpha^{-\frac{5}{2}} t^{-\frac{1}{2}} + 3\alpha^{-\frac{3}{2}} t^{-\frac{1}{2}} \right) \\ &\leq \sqrt{G} \sum_{t=1}^{T} \left(2\alpha^{-\frac{5}{2}} t^{-\frac{1}{2}} + 3\alpha^{-\frac{3}{2}} t^{-\frac{1}{2}} \right) \\ &\leq O\left(\left(\alpha^{-\frac{5}{2}} + \alpha^{-\frac{3}{2}} \right) \sqrt{GT} \right) \end{split}$$
(16)

Therefore, we conclude the proof.

B.3 PROOF OF LEMMA 5

Proof. For convenience, we define $\tilde{k}_y := \tilde{\Phi}_g(\varphi(y)) \in \mathbb{R}^{|\Psi_g|}, Q := (K + \alpha I)^{-1} \in \mathbb{R}^{|\Psi_g| \times |\Psi_g|}$, and $\widetilde{Q} := (\widetilde{K} + \alpha I)^{-1} \in \mathbb{R}^{|\Psi_g| \times |\Psi_g|}$, where $\widetilde{K} := \widetilde{\Phi}_g \widetilde{\Phi}_g^{\top}$. Using the same notations in the proof of Lemma 1, we obtain that

$$\begin{aligned} |\widetilde{\mu}_{g} - \widehat{\mu}_{g}| &= \left| (\varphi(y))^{\top} \widetilde{\Phi}_{g}^{\top} (\widetilde{K} + \alpha I)^{-1} v - k_{y}^{\top} (K + \alpha I)^{-1} v \right| \\ &= \left| \widetilde{k}_{y}^{\top} \widetilde{Q} v - k_{y}^{\top} Q v \right| \\ &\leq \left| \widetilde{k}_{y}^{\top} (\widetilde{Q} - Q) v \right| + \left| (\widetilde{k}_{y} - k_{y})^{\top} Q v \right|, \end{aligned}$$

$$(17)$$

where we use Equation (22) to derive $\widetilde{\mu}_g = (\varphi(y))^\top (\widetilde{\Phi}_g^\top \widetilde{\Phi}_g + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top (\widetilde{K} + \alpha I)^{-1} \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_g^\top v = (\varphi(y))^\top \widetilde{\Phi}_$ $\alpha I)^{-1}v$ in the first equation.

1. Bounding $|(\tilde{k}_y - k_y)^\top Qv|$. We evoke (Rahimi & Recht, 2007a, Claim 1), which is rewritten in Lemma 8 using our notations. For a desired threshold $\epsilon_{\text{RFF}} > 0$, set

$$s = \frac{4(d+2)}{\epsilon_{\rm RFF}^2} \log \left(\frac{\sigma_p^2}{(\delta/2) \cdot \epsilon_{\rm RFF}^2} \cdot 2^8 \right).$$

Then, with probability at least $1 - \frac{\delta}{2}$, it holds that $\sup_{y,y' \in \mathcal{Y}} |(\varphi(y))^\top \varphi(y') - k(y,y')| \le \epsilon_{\text{RFF}}$, and hence $\|\tilde{k}_y - k_y\|_{\infty} \le \epsilon_{\text{RFF}}$. Therefore, we obtain

$$|(\tilde{k}_y - k_y)^{\top} Q v| \le \|\tilde{k}_y - k_y\|_2 \cdot \|Q\|_2 \cdot \|v\|_2 \le \epsilon_{\text{RFF}} \sqrt{|\Psi_g|} \cdot \alpha^{-1} \cdot \sqrt{|\Psi_g|} = \alpha^{-1} |\Psi_g| \epsilon_{\text{RFF}},$$
(18)

where the last inequality holds by $||Q||_2 = \lambda_{\min}^{-1}(K + \alpha I) \le \alpha^{-1}$ and $||v||_{\infty} \le 1$.

2. Bounding $|\tilde{k}_y^{ op}(\widetilde{Q}-Q)v|$. Note that

$$\begin{split} |\tilde{k}_y^{\top}(\widetilde{Q} - Q)v| \leq & \|\tilde{k}_y\|_2 \cdot \|\widetilde{Q} - Q\|_2 \cdot \|v\|_2 \\ \leq & \sqrt{2|\Psi_g|} \cdot \|\widetilde{Q} - Q\|_2 \cdot \sqrt{|\Psi_g|} \end{split}$$

where the first inequality holds by the fact that $(\varphi(y))^{\top}\varphi(y_i) = (2/s)\sum_{j=1}^s \cos(w_j^{\top}y + b_j)\cos(w_{i,j}^{\top} + b_{i,j}) \leq 2$. To bound $\|\tilde{Q} - Q\|_2$, we evoke (Avron et al., 2017, Theorem 7), which is rewritten in Lemma 9. For a desired threshold $\Delta_{\text{RFF}} \leq 1/2$, the following inequality holds with probability at least $1 - \frac{\delta}{2}$:

$$(1 - \Delta_{\text{RFF}})(K + \alpha I) \preceq \widetilde{K} + \alpha I$$

for $s \ge \frac{8|\Psi_g|}{3\alpha} \Delta_{\text{RFF}}^{-2} \log(32s_\alpha(K)/\delta)$. By Sherman-Morrison-Woodbury formula, i.e., $A^{-1} - B^{-1} = A^{-1}(B-A)B^{-1}$ where A and B are invertible, we derive

$$\begin{aligned} &\|(\widetilde{K} + \alpha I)^{-1} - (K + \alpha I)^{-1}\|_{2} \\ &\leq \|(\widetilde{K} + \alpha I)^{-1}\|_{2} \cdot \|(K + \alpha I) - (\widetilde{K} + \alpha I)\|_{2} \cdot \|(K + \alpha I)^{-1}\|_{2} \end{aligned}$$

$$< \alpha^{-2} \Delta_{\text{RFF}}(\|K\|_2 + \alpha)$$

where the last inequality holds by the fact that $\|(\widetilde{K} + \alpha I)^{-1}\|_2$, $\|(K + \alpha I)^{-1}\|_2 \le \alpha^{-1}$ and $\|(K + \alpha I) - (\widetilde{K} + \alpha I)\|_2 \le \|\Delta_{\text{RFF}}(K + \alpha I)\|_2 \le \Delta_{\text{RFF}}(\|K\|_2 + \alpha)$.

3. Putting everything together. Combining Equations (18) and (19), with probability at least $1 - \delta$, it holds that

$$|\widetilde{\mu}_g - \widehat{\mu}_g| \le \alpha^{-1} |\Psi_g| \epsilon_{\text{RFF}} + \alpha^{-2} \Delta_{\text{RFF}} (||K||_2 + \alpha)$$

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$$s \ge \max\left\{\frac{4(d+2)}{\epsilon_{\rm RFF}^2}\log\left(\frac{\sigma_p^2}{(\delta/2)\cdot\epsilon_{\rm RFF}^2}\cdot 2^8\right), \frac{8|\Psi_g|}{3\alpha}\Delta_{\rm RFF}^{-2}\log\left(\frac{32s_\alpha(K)}{\delta}\right)\right\},$$
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which concludes the proof.

(19)

B.4 PROOF OF LEMMA 6

 $|\widetilde{\sigma}_a - \widehat{\sigma}_a|$

Proof. We use the same notations in the proof of Lemma 5. Let c denote a lower bound of $1 - ||k_y||^2_{(K+\alpha I)^{-1}}$. Note that

$$= \alpha^{-\frac{1}{2}} \left| \sqrt{1 - (\varphi(y))^{\top} \tilde{\Phi}_{g}^{\top} (\tilde{K} + \alpha I)^{-1} \tilde{\Phi}_{g}(\varphi(y))} - \sqrt{1 - k_{y}^{\top} (K + \alpha I)^{-1} k_{y}} \right|$$

$$\leq (c \cdot \alpha)^{-\frac{1}{2}} \left| (\varphi(y))^{\top} \tilde{\Phi}_{g}^{\top} (\tilde{K} + \alpha I)^{-1} \tilde{\Phi}_{g}(\varphi(y)) - k_{y}^{\top} (K + \alpha I)^{-1} k_{y} \right|$$

$$= (c \cdot \alpha)^{-\frac{1}{2}} \left| \tilde{k}_{y}^{\top} \widetilde{Q} \tilde{k}_{y} - k_{y}^{\top} Q k_{y} \right|$$

$$\leq (c \cdot \alpha)^{-\frac{1}{2}} \left(\left| \tilde{k}_{y}^{\top} (\widetilde{Q} - Q) \tilde{k}_{y} \right| + \left| (\tilde{k}_{y} - k_{y})^{\top} Q \tilde{k}_{y} \right| + \left| k_{y}^{\top} Q (\tilde{k}_{y} - k_{y}) \right| \right)$$

$$\leq (c \cdot \alpha)^{-\frac{1}{2}} \left(\left\| \tilde{k}_{y} \right\|_{2}^{2} \| \widetilde{Q} - Q \|_{2} + \| \tilde{k}_{y} - k_{y} \|_{2} \| \tilde{k}_{y} \|_{2} \| Q \|_{2} + \| \tilde{k}_{y} - k_{y} \|_{2} \| k_{y} \|_{2} \| Q \|_{2} \right)$$

$$\leq (c \cdot \alpha)^{-\frac{1}{2}} \left(2 |\Psi_{g}| \alpha^{-2} \Delta_{\text{RFF}} (\| K \|_{2} + \alpha) + 3\alpha^{-1} |\Psi_{g}| \epsilon_{\text{RFF}} \right)$$
(20)

which concludes the proof.

B.5 ANALYSIS OF LEMMA 2

Proof. Solving KRR with n regression data requires $\Theta(n^3)$ time and $\Theta(n^2)$ space. Hence, by the convexity of the cubic and quadratic functions, the time for COMPUTE_UCB scales with $\Theta(\sum_{g\in\mathcal{G}}n_g^3) = O(t^3/G^2)$, and the space scales with $\Theta(\sum_{g\in\mathcal{G}}n_g^2) = O(t^2/G)$, where $n_g := |\Psi_g|$ is the visitation to any model $g \in \mathcal{G}$ up to iteration t, and we have $\sum_{g \in \mathcal{G}} n_g = t$. On the other hand, solving KRR with n regression data and random features of size s requires $O(ns^2)$ time and O(ns)space. Therefore, the time for COMPUTE_UCB_RFF scales with $O(\sum_{q \in \mathcal{G}} n_g s^2) = O(ts^2)$, and the space scales with $O(\sum_{g \in \mathcal{G}} n_g s) = O(ts)$.

1111 C AUXILIARY LEMMAS

1113 C.1 PROOF OF LEMMA 1

Proof. We rewrite the proof using the notations in Section 5. Obviously, Equation (5) holds when 1116 the index set Ψ_g is empty. In the following, we consider non-empty Ψ_g . Let $\Phi_g := [\phi(y_i)^\top]_{i \in \Psi_g}$. 1117 Note that $k_y = [k(y, y_i)]_{i \in \Psi_g}^\top = \Phi_g(\phi(y))$ and $K = [k(y_i, y_j)]_{i,j \in \Psi_g} = \Phi_g \Phi_g^\top$. We have

$$\begin{aligned} \widehat{\mu}_{g} - s_{g}(y) &= (\phi(y))^{\top} \Phi_{g}^{\top} (K + \alpha I)^{-1} v - (\phi(y))^{\top} w_{g}^{\star} \\ &= (\phi(y))^{\top} (\Phi_{g}^{\top} \Phi_{g} + \alpha I)^{-1} \Phi_{g}^{\top} v - (\phi(y))^{\top} (\Phi_{g}^{\top} \Phi_{g} + \alpha I)^{-1} (\Phi_{g}^{\top} \Phi_{g} + \alpha I) w_{g}^{\star} \end{aligned} (21) \\ &= (\phi(y))^{\top} (\Phi_{g}^{\top} \Phi_{g} + \alpha I)^{-1} \Phi_{g}^{\top} (v - \Phi_{g} w_{g}^{\star}) - \alpha(\phi(y))^{\top} (\Phi_{g}^{\top} \Phi_{g} + \alpha I)^{-1} w_{g}^{\star}, \end{aligned}$$

where the second equation holds by the positive definiteness of both matrices $(K+\alpha I)$ and $(\Phi_g^{\top}\Phi_g + \alpha I)$ and hence

$$\Phi_{g}^{\top}(K + \alpha I)^{-1} = (\Phi_{g}^{\top}\Phi_{g} + \alpha I)^{-1}\Phi_{g}^{\top}.$$
(22)

1. Bounding $(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} \Phi_g^{\top} (v - \Phi_g w_g^{\star})$. Note that the scores $\{s_t : t \in \Psi_g\}$ are 1130 independent by the construction of Φ_g and $\mathbb{E}[s_t] = (w_g^{\star})^{\top} \phi(y_t)$, we have that

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$$(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} \Phi_g^{\top} (v - \Phi_g w_g^{\star}) = \sum_{i=1}^{|\Psi_g|} [(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} \Phi_g^{\top}]_i \cdot [v - \Phi_g w_g^{\star}]_i$$

are summation of zero mean independent random variables, where we denote by $[\cdot]_i$ the *i*-th element of a vector. Further, each variable satisfies that

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$$\left| \left[(\phi(y))^\top (\Phi_g^\top \Phi_g + \alpha I)^{-1} \Phi_g^\top \right]_i \cdot \left[v - \Phi_g w_g^\star \right]_i \right| \right|$$

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$$\leq \|(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} \Phi_g^{\top} \| \cdot |[v - \Phi_g w_g^{\star}]_i|$$
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$$\leq \sqrt{(\phi(y))^\top (\Phi_g^\top \Phi_g + \alpha I)^{-1} \Phi_g^\top \Phi_g (\Phi_g^\top \Phi_g + \alpha I)^{-1} (\phi(y))} \cdot (1 + \|w_g^\star\|)$$

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1143 where the last inequality holds by $||w_a^*|| \le 1$ and the second inequality holds by

1152 1153 Then, by Azuma-Hoeffding inequality, it holds that

 $\leq 2\widehat{\sigma}_a$

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$$\mathbb{P}\left(|(\phi(y))^{\top}(\Phi_g^{\top}\Phi_g + \alpha I)^{-1}\Phi_g^{\top}(v - \Phi_g w_g^{\star})| > 2\eta\widehat{\sigma}_g\right)$$

$$\leq 2\exp\left(-\frac{\widehat{\sigma}_g^2\eta^2}{2|\Psi_g|\widehat{\sigma}_g^2}\right)$$
(23)

1160 1161 1. Bounding $\alpha(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} w_g^{\star}$. By the Cauchy-Schwarz inequality, it holds that **1162 1163** $|(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} w_g^{\star}|$ **1164 1165** $(\phi(y))^{\top} (\Phi_g^{\top} \Phi_g + \alpha I)^{-1} w_g^{\star}|$

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$$\leq \|w_g\| \cdot \|(\phi(y))^\top (\Phi_g^\top \Phi_g + \alpha I)^- \|$$

$$= \|w_g^\star\| \cdot \sqrt{(\phi(y))^\top (\Phi_g^\top \Phi_g + \alpha I)^{-1} \alpha^{-1} \alpha I (\Phi_g^\top \Phi_g + \alpha I)^{-1} (\phi(y))}$$

$$\leq \alpha^{-1/2} \sqrt{(\phi(y))^\top (\Phi_g^\top \Phi_g + \alpha I)^{-1} (\Phi_g^\top \Phi_g + \alpha I) (\Phi_g^\top \Phi_g + \alpha I)^{-1} (\phi(y))}$$

$$= \alpha^{-1/2} \widehat{\sigma}_g,$$

(24)

where the second inequality holds by the positive definiteness of $\Phi_{g}^{\top}\Phi_{g}$.

3. Putting everything together. Plugging (23) and (24) in (21) and setting $\delta = 2 \exp(-\eta^2/2)$ concludes the proof.

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1176 C.2 PROOF OF LEMMA 4

1178 *Proof.* The first statement holds by both Lemma 3 and Lemma 1, and a uniform bound over all 1179 $t \in [T], g \in \mathcal{G}$, and $m \in [M]$. Let $g_{\star,t} := \arg \max_{g \in \mathcal{G}} s_g(y_t)$ is the optimal model for prompt y_t 1180 and $\widehat{g}_{\star,t}^m := \arg \max_{g \in \widehat{\mathcal{G}}_{\star}^m} \widehat{s}_{g,t}^m$ is optimistic model at stage m.

1181 1182 To show the second statement, first note that $g_{\star,t} \in \widehat{\mathcal{G}}_t^1$. Assume $g_{\star,t} \in \widehat{\mathcal{G}}_t^m$ for some $m \in [M-1]$. 1183 Then, by the first statement, we obtain that $\widehat{s}_{g_{\star,t},t}^m - \max_{g \in \widehat{\mathcal{G}}_t^m} \widehat{s}_{g,t}^m \ge s_{g_{\star,t}}(y_t) - (2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_{g_{\star,t}}^m - (s_{\widehat{g}_{\star,t}^m}(y_t) + (2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_{\widehat{g}_{\star,t}^m}^m) \ge -(2\eta + \sqrt{\alpha}) \cdot (\widehat{\sigma}_{g_{\star,t}}^m + \widehat{\sigma}_{\widehat{g}_{\star,t}^m}^m) \ge 2 \cdot 2^{-m} = 2^{1-m}$, which ensures 1185 $g_{\star,t} \in \widehat{\mathcal{G}}_t^{m+1}$.

Finally, by the first two statements, we have that $s_{\star}(y_t) - s_g(y_t) \leq \widehat{s}_{g_{\star,t},t}^m + 2(2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_{g_{\star,t}}^m - (\widehat{s}_{g,t}^m - 2(2\eta + \sqrt{\alpha}) \cdot \widehat{\sigma}_g^{m,t}) \leq 2 \cdot 2^{1-m} = 2^{3-m}$. We conclude the proof.

1188 C.3 USEFUL LEMMAS

Lemma 8 (Rahimi & Recht (2007a), Claim 1). Let \mathcal{M} be a compact subset of \mathbb{R}^d with diameter diam(\mathcal{M}). Then, for the mapping $\varphi : \mathbb{R}^d \to \mathbb{R}^s$ defined in Equation (3), we have

1193 1194 $\mathbb{P}\left(\sup_{y,y'\in\mathcal{M}} |(\varphi(y))^{\top}\varphi(y') - k(y,y')| \ge \epsilon\right) \le 2^{8} \left(\frac{\sigma_{p} \cdot \operatorname{diam}(\mathcal{M})}{\epsilon}\right)^{2} \exp\left(-\frac{s\epsilon^{2}}{4(d+2)}\right),$

1195 where $\sigma_p^2 := \mathbb{E}_p[\omega^{\top}\omega]$ is the second moment of the Fourier transform of k.⁷ Further, 1196 $\sup_{y,y'\in\mathcal{M}} |(\varphi(y))^{\top}\varphi(y') - k(y,y')| \le \epsilon$ with any constant probability when $s = \Omega((d/\epsilon^2)\log(\sigma_p \cdot diam(\mathcal{M})/\epsilon))$.

1198 1199 1199 1200 1201 1202 Lemma 9 (Avron et al. (2017), Theorem 7). Let $K = [k(y_i, y_j)]_{i,j \in [n]}$ denote the Gram matrix of $\{y_i \in \mathbb{R}^d\}_{i=1}^n$, where k is a shift-invariant kernel function. Let $\Delta \leq 1/2$ and $\delta \in (0, 1)$. Assume that $\|K\|_2 \geq \alpha$. If we use $s \geq \frac{8n}{3\alpha} \Delta^{-2} \log(16s_\alpha(K)/\delta)$ random Fourier features, then with probability at least $1 - \delta$, it holds that

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 $(1 - \Delta)(K + \alpha I) \preceq \widetilde{K} + \alpha I \preceq (1 + \Delta)(K + \alpha I)$

where $s_{\alpha}(K) := \text{Tr}[(K + \alpha I)^{-1}K]$ and we denote by $\widetilde{K} = [(\varphi(y_i))^{\top}(\varphi(y_j))]_{i,j\in[n]}$ the approximated Gram matrix using $s \in \mathbb{N}_+$ random Fourier features, where $\varphi : \mathbb{R}^d \to \mathbb{R}^s$ is the feature mapping.

Lemma 10 (Valko et al. (2013), Lemma 4). For any model $g \in \mathcal{G}$ and stage $m \in [M]$, let $\lambda_{g,1}^m \geq \lambda_{g,2}^m \geq \cdots$ denote the eigenvalues (in the decreasing order) of the matrix $(\Phi_g^m)^\top \Phi_g^m + \alpha I$, where $\Phi_g^m = [\phi(y_i)^\top]_{i \in \Psi_{g,T+1}^m}$. Then, for any iteration $t \in [T]$, it holds that

$$\sum_{t \in \Psi_{g,T+1}^m} \widehat{\sigma}_{g,t}^m \leq \widetilde{O}\left(\sqrt{d_{\text{eff}}\left(10 + \frac{15}{\alpha}\right)|\Psi_{g,T+1}^m|}\right)$$

where $d_{\text{eff}} := \max_{g \in \mathcal{G}, m \in [M]} \min\{j \in \mathbb{N}_+ : j\alpha \log T \ge \Lambda_{g,j}^m\}$ and $\Lambda_{g,j}^m := \sum_{i>j} \lambda_{g,i}^m - \alpha$ is the effective dimension.

D ADDITIONAL EXPERIMENTAL DETAILS AND RESULTS

1. Implementation details. We use the CLIP-based features of the prompts as the context vector (Cherti et al., 2023) for tasks of T2I and T2V generation, and we use the CLIP-based features of the images in the task of image captioning. We set both the exploration and regularization parameters $\alpha, \eta = 1$ in all the experiments. Two hyperparameters have to be chosen. The first one is the parameter γ in the polynomial and radial basis function (RBF) kernels, which are given by

$$k_{\text{poly3}}^{\gamma}(x_1, x_2) = (\gamma \cdot x_1^{\top} x_2 + 1)^3, \quad k_{\text{RBF}}^{\gamma}(x_1, x_2) = \exp(-\gamma \cdot \|x_1 - x_2\|^2).$$

1226 In the experiments, we select γ to be 5 and 1 for the polynomial and RBF kernel functions, respec-1227 tively. The second hyperparameter is the number of random features in the RFF-UCB algorithm. In 1228 addition, the features are generated once for each sample and stored to save the computation of the 1229 RFF-UCB algorithm.

2. Ablation study on hyperparameters. We conduct ablation studies on the selections of pa-1231 rameter γ in the RBF kernel function and the number of features in RFF-UCB. The results are 1232 summarized in Figures 8 and 9, respectively. We select $\gamma = 1$ (default), 3, 5, and 7 and num-1233 ber of features varying between 25, 50 (default), 75, and 100. Results show that the RFF-UCB 1234 algorithm can attain consistent performance. [Additionally, we test the SCK-UCB-poly3 algo-1235 rithm with $\gamma = 1, 3, 5$ (default), and 7 in the polynomial kernel and regularization parameter 1236 $\alpha = 0.5, 1.0$ (default), and 1.5 in KRR. The results are summarized in Figures 10 and 11, re-1237 spectively.] 1238

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⁷For the RBF kernel with parameter σ^2 , i.e., $k_{\text{RBF}}(y, y') = \exp(-\frac{1}{2\sigma^2} \|y - y'\|_2^2)$, we have $\sigma_{p_{\text{RBF}}}^2 = \frac{d}{\sigma^2}$.







4. Additional examples. We provide more examples showing that prompt-based generative models can outperform for text prompts from certain categories while underperforming for other text categories (see Figures 17 and 18).





Figure 17: Prompt-based generated images from uniDiffuser (Bao et al., 2023) and PixArt- α : uni-Diffuser attains a higher CLIPScore in generating type "train" prompts (35.29 versus 34.25) while underperforms for type "baseball bat" prompts (32.51 versus 34.30).

