EARTHQUAKENPP: BENCHMARK DATASETS FOR EARTHQUAKE FORECASTING WITH NEURAL POINT PROCESSES

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ABSTRACT

Classical point process models, such as the epidemic-type aftershock sequence (ETAS) model, have been widely used for forecasting the event times and locations of earthquakes for decades. Recent advances have led to Neural Point Processes (NPPs), which promise greater flexibility and improvements over classical models. However, the currently-used benchmark dataset for NPPs does not represent an up-to-date challenge in the seismological community since it lacks a key earthquake sequence from the region and improperly splits training and testing data. Furthermore, initial earthquake forecast benchmarking lacks a comparison to state-of-the-art earthquake forecasting models typically used by the seismological community. To address these gaps, we introduce EarthquakeNPP: a collection of benchmark datasets to facilitate testing of NPPs on earthquake data, accompanied by a credible implementation of the ETAS model. The datasets cover a range of small to large target regions within California, dating from 1971 to 2021, and include different methodologies for dataset generation. In a benchmarking experiment, we compare three spatio-temporal NPPs against ETAS and find that none outperform ETAS in either spatial or temporal log-likelihood. These results indicate that current NPP implementations are not yet suitable for practical earthquake forecasting. EarthquakeNPP also provides generative evaluation metrics, enabling broader model classes to be benchmarked and facilitating the future collaboration between the seismology and machine learning communities.

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1 INTRODUCTION

Operational earthquake forecasting by global governmental organisations such as the US Geological Survey (USGS) necessitates the development of models which can forecast the times and locations of damaging earthquakes. While model development is ongoing in the seismology community, recent improvements have relied upon refinement of a spatio-temporal point process model known as the Epidemic-Type Aftershock Sequence (ETAS) model (Ogata, 1988; 1998), despite significant growth in available data (Takanami et al., 2003; Shelly, 2017; Ross et al., 2019; White et al., 2019; Mousavi et al., 2020; Tan et al., 2021; Mousavi & Beroza, 2023).

In contrast, the machine learning community has offered promising advancements over classical point process models like ETAS with Neural Point Process (NPP) models, showcasing greater flexibility (Du et al., 2016; Omi et al., 2019a; Shchur et al., 2019; Jia & Benson, 2019; Chen et al., 2021; Zhou et al., 2022; Zhou & Yu, 2024). While some initial benchmarking of these models has been conducted on an earthquake dataset in Japan, these experiments lack relevance for stakeholders in the seismology community. The benchmark lacks a key earthquake sequence from the region, fails to recreate an operational setting with proper train-test splits, and doesn't compare against state-of-the-art models like ETAS.

Here, we introduce EarthquakeNPP: a curated collection of datasets designed for benchmarking
 NPP models in earthquake forecasting, accompanied by a state-of-the-art benchmark model. These
 datasets are derived from publicly available raw data, which we process and configure within our
 platform to facilitate meaningful forecasting experiments relevant to stakeholders in the seismology
 community. Covering various regions of California, these datasets represent typical forecasting zones

Table 1: Comparison of EarthquakeNPP datasets with the existing NPP benchmark dataset for
 earthquakes.

Dataset	Chronological Training/Test Splits	Complete Timespan	Complete Magnitudes	Used by Local Agencies
Chen et al. (2021) Dataset	×	×	×	×
EarthquakeNPP Datasets	 Image: A second s	 Image: A second s	 Image: A second s	 Image: A second s

and encompass data commonly utilized by forecast issuers. Moreover, employing modern techniques, some datasets include smaller magnitude earthquakes, exploring the potential of numerous small events to enhance forecasting performance through flexible NPPs. To unify efforts, we present an operational-level implementation of the ETAS model alongside the datasets, serving as a benchmark for NPPs.

Although initial benchmarking finds that none of the 3 tested NPP implementations outperform ETAS, EarthquakeNPP aims to serve as a platform for future NPP development. The platform facilitates the generative evaluation procedure used for rigorous benchmarking in the seismology community. This directs the impact of future NPPs to stakeholders in seismology and broadens the scope of models beyond NPPs (e.g. times series models (Wang et al., 2017), Bayesian approaches (Serafini et al., 2023)). Access to the dataset collection, along with comprehensive documentation and notebooks, can be found at https://anonymous.4open.science/r/EarthquakeNPP-2D51.

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1.1 RELATED WORK

077 Benchmarking by the NPP Community. Chen et al. (2021) introduced an earthquake dataset for 078 benchmarking the Neural Spatio-temporal Point Process (NSTPP) model using a global dataset from 079 the U.S. Geological Survey, focusing on Japan from 1990 to 2020. They considered earthquakes with magnitudes above 2.5, splitting the data into month-long segments with a 7-day offset. They exclude earthquakes from November 2010 to December 2011, deeming these sequences "too long" 081 and "outliers." However, this period includes the 2011 Tohoku earthquake (Mori et al., 2011), the largest earthquake recorded in Japan and the fourth largest in the world, at magnitude 9.0. This 083 exclusion renders the benchmarking experiment irrelevant for seismologists, as it is precisely these 084 large earthquakes and their aftershocks that are crucial to forecast due to their damaging impact. 085 Additionally, these events are of significant scientific interest because they provide valuable insights into the earthquake rupture process. 087

The dataset segments are divided for training, testing, and validation. Instead of a chronological partitioning that mirrors operational forecasting, the segments are assigned in an alternating pattern. This approach misrepresents a realistic forecasting scenario and inflates performance measures due to earthquake triggering (Freed, 2005). Since the model is tested on windows immediately preceding training windows, it exploits causal dependencies backwards it time.

Although earthquakes with magnitudes above 2.5 are considered by Chen et al. (2021), following a change in USGS policy on global data collection, from 2009 onwards, only events above magnitude
4.0 are recorded in the dataset. For earthquake forecasting in Japan, seismologists use datasets from Japanese data centers since they are more comprehensive and complete than global datasets. Section A.2 describes the biases incurred from such data missingness.

Chen et al. (2021) benchmark their model against another spatio-temporal model, Neural Jump SDEs (Jia & Benson, 2019), and a temporal-only Hawkes process, even though a spatio-temporal Hawkes process would provide a more rigorous benchmark. Subsequent papers adopting this benchmark (Zhou et al., 2022; Yuan et al., 2023; Zhou & Yu, 2024) similarly lack comparisons to a spatio-temporal Hawkes process, benchmarking instead against temporal-only or spatial-only baselines or other spatio-temporal NPPs.

Benchmarking by the Seismology Community. Model comparison has been crucial in the develop ment of earthquake forecasting models since their inception (Kagan & Knopoff, 1987; Ogata, 1988).
 The Collaboratory for the Study of Earthquake Predictability (CSEP) (Michael & Werner, 2018; Schor lemmer et al., 2018; Savran et al., 2022; Iturrieta et al., 2024) (https://cseptesting.org/) aims to unify the framework for earthquake model testing and evaluation, hosting retrospective

and fully prospective forecasting experiments globally. CSEP benchmarks short-term models using
 performance metrics that require forecasts to be generated by simulating many repeat sequences over
 a specified time horizon (typically one day). These simulated forecasts are compared by discretizing
 time and space intervals, with test statistics calculated for event counts, magnitudes, locations, and
 times. The simulation-based approach allows the inclusion of generative models that don't output
 explicit earthquake probabilities (i.e., a likelihood), and enables evaluation of the full distribution of
 entire sampled sequences.

Two existing works benchmark NPPs for earthquake forecasting within the seismology community. The first by Dascher-Cousineau et al. (2023) extends a temporal-only NPP from Shchur et al. (2019) to include earthquake magnitudes. The second by Stockman et al. (2023) extends another temporal-only model by Omi et al. (2019a) to target larger magnitude events. Both models are benchmarked against a temporal ETAS model, showing moderate improvements over the baseline. Extending these models to include spatial data is necessary for further testing and potential operational use in the seismological community.

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1.2 Scope of this work

Since generating repeated sequences over forecast horizons is computationally costly, the seismology community uses the mean log-likelihood on held-out data for a more streamlined metric during model development (Ogata, 1988; Harte, 2015). Our platform uses this metric in the NPP benchmarking experiment and provides detailed guidance on CSEP's simulation-based procedure, enabling future NPP implementations and evaluations within CSEP experiments.

129 This work aims to bridge Machine Learning and seismology by establishing a baseline for com-130 paring NPP models to state-of-the-art, domain-based models. Only NPPs capable of generating 131 log-likelihoods are within scope, as no valid score exists for models lacking this capability (e.g. 132 Yuan et al., 2023; Li et al., 2023). Traditional metrics like Root Mean Square Error (RMSE) and 133 Mean Absolute Error (MAE) are inadequate and potentially misleading for seismological predictions 134 (Hodson, 2022), as earthquake occurrence follows power law distributions (Kagan, 1994; Felzer & 135 Brodsky, 2006) that are heavy-tailed, making the errors non-Gaussian and non-Laplacian — contrary 136 to the assumptions underlying RMSE and MAE (see Section G). To ensure seismological relevance, we challenge authors of NPP models to implement forecasts using CSEP's evaluation framework and 137 benchmark their results against the performance of the ETAS model. 138

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2 BACKGROUND

2.1 Spatio-Temporal Point Processes

144 A spatio-temporal point process is a continuous-time stochastic process that models the random 145 number of events $N(S \times (t_a, t_b])$ which occur in a space-time interval $S \times (t_a, t_b]$, $S \in \mathbb{R}^2$, $(t_a, t_b] \in \mathbb{R}^+$. This process is typically defined by a non-negative *conditional intensity function*

$$\lambda(t, \mathbf{x} | \mathcal{H}_t) := \lim_{\Delta t, \Delta \mathbf{x} \to 0} \frac{\mathbb{E}\left[N([t, t + \Delta t) \times B(\mathbf{x}, \Delta \mathbf{x}) | \mathcal{H}_t\right]}{|B(\mathbf{x}, \Delta \mathbf{x})|},\tag{1}$$

where $\mathcal{H}_t = \{(t_i, \mathbf{x}_i) | t_i < t\}$ denotes the history of events preceding time t and $|B(\mathbf{x}, \Delta \mathbf{x})|$ is the Lebesgue measure of the ball $B(\mathbf{x}, \Delta \mathbf{x})$ with radius $\Delta \mathbf{x}$. Given we observe a history of events up to t_i , the probability density function (pdf) of observing an event at time t and location \mathbf{x} is given by

$$p(t, \mathbf{x} | \mathcal{H}_{t_i}) = \lambda(t, \mathbf{x} | \mathcal{H}_{t_i}) \cdot \exp\left(-\int_{t_i}^t \int_{\mathcal{S}} \lambda(s, \mathbf{z} | \mathcal{H}_s) d\mathbf{z} ds\right).$$
 (2)

Most models specify the conditional intensity function, though some (e.g. Shchur et al., 2019; Chen et al., 2021; Yuan et al., 2023) directly model this pdf. Model parameters are typically estimated by maximizing the log-likelihood of observed events within a training time interval $[T_0, T_1]$ and spatial region S,

$$\log p(\mathcal{H}_T) = \underbrace{\sum_{i=0}^{n} \log \lambda(t_i | \mathcal{H}_{t_i}) - \int_{T_0}^{T_1} \int_{\mathcal{S}} \lambda(s, \mathbf{z} | \mathcal{H}_s) d\mathbf{z} ds}_{\mathcal{S}} + \underbrace{\sum_{i=0}^{n} \log f(\mathbf{x}_i | t_i, \mathcal{H}_{t_i})}_{\mathcal{S}}, \quad (3)$$

Spatial log-likelihood

where the decomposition of the spatio-temporal conditional intensity function, $\lambda(t_i, \mathbf{x}_i | \mathcal{H}_{t_i}) = \lambda(t_i | \mathcal{H}_{t_i}) \cdot f(\mathbf{x}_i | t_i, \mathcal{H}_{t_i})$, allows the log-likelihood to be written as contributions from the temporal and spatial components. In practice, this exact function is often not maximized directly during training: for models specified through the conditional intensity function, an analytical solution to the integral term is generally not possible and is approximated numerically.

167 For model evaluation and comparison, the log-likelihood of observing events in the test set can 168 be used as a performance metric. This is consistent with a wealth of literature in the seismology 169 community (see Zechar et al., 2010, and references therein) as well as the wider general point process 170 literature (Daley & Vere-Jones, 2004), which now includes neural point processes (Shchur et al., 171 2021). The metric evaluates models that output probability distributions over their predictions and 172 consequently penalises models that are overconfident. Although evaluating on events in the test set, the test log-likelihood, $\log p((t_i, \mathbf{x}_i) | t_i \in [T_2, T_3], \mathcal{H}_{T_2})$, may still contain dependence upon events 173 prior to the test window $[T_2, T_3]$, typically contained in the history \mathcal{H}_{T_2} of the intensity function. 174 Comparing the mean log-likelihood per event provides the *information gain* from one model to 175 another (Daley & Vere-Jones, 2004). 176

177 Point processes are the dominant modeling approach in the seismology community, used exten-178 sively in both real-time operational earthquake forecasting (Mizrahi et al., 2024a) and established 179 benchmarking experiments (CSEP) (Taroni et al., 2018; Rhoades et al., 2018). The point process representation of earthquake data aligns naturally with their occurrence as discrete events in time 180 (Kagan, 1994). Furthermore, this modeling approach is favored over discretized forecasting models 181 (e.g., time series) because it eliminates the need for optimizing binning strategies and allows for 182 immediate updates, rather than waiting until the end of a time bin — a delay that could miss critical, 183 potentially damaging events. 184

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207 208 2.2 ETAS

The Epidemic Type Aftershock Sequence (ETAS) model (Ogata, 1998) is a spatio-temporal Hawkes process which models how earthquakes cluster in time and space. It has been adopted for operational earthquake forecasting by government agencies in California (Milner et al., 2020), New-Zealand (Christophersen et al., 2017), Italy (Spassiani et al., 2023), Japan (Omi et al., 2019b) and Switzerland (Mizrahi et al., 2024b), and performs consistently well in CSEP's retrospective and fully prospective forecasting experiments (e.g. Woessner et al., 2011; Rhoades et al., 2018; Taroni et al., 2018; Cattania et al., 2018; Mancini et al., 2019; 2020; 2022). The general formulation of the model is

$$\lambda(t, \mathbf{x} | \mathcal{H}_t; \theta) = \mu + \sum_{i: t_i < t} g(t - t_i, ||\mathbf{x} - \mathbf{x}_i||_2^2, m_i),$$
(4)

where μ is a constant background rate of events, $g(\cdot, \cdot, \cdot)$ is a non-negative excitation kernel which describes how past events contribute to the likelihood of future events and m_i are the associated magnitudes of each event. The equivalent formulation as a Hawkes branching process accompanies a causal branching structure **B**. This concept broadly aligns with the understanding of the physics of earthquake triggering and interaction, e.g. via dynamic wave triggering (Brodsky & van der Elst, 2014) and static stress triggering (Gomberg, 2018; Mancini et al., 2020).

Although ETAS can be fit by maximizing the log-likelihood function directly, parameter estimation is typically performed by simultaneously estimating the branching structure **B**. Veen & Schoenberg (2008) developed an Expectation Maximisation (EM) procedure, which maximises the marginal likelihood over the unobserved branching structure, $\log \int p(\mathcal{H}_{T_1}|\mathbf{B}, \theta) p(\mathbf{B}|\theta) d\mathbf{B}$ through the iteration

$$\theta^{(k+1)} = \arg\max_{\theta} \mathbb{E}_{\mathbf{B} \sim p(\cdot | \mathcal{H}_{T_1}, \theta^{(k)})} \left[\log p(\mathcal{H}_{T_1}, \mathbf{B} | \theta) \right].$$
(5)

This avoids the need to numerically approximate the integral term in the likelihood, provides more stability during estimation and simultaneously distinguishes background events from triggered events.

The formulation of the ETAS model we present with the EarthquakeNPP datasets is implemented in the etas python package by Mizrahi et al. (2022). It defines the triggering kernel as

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$$g(t, r^2, m) = \frac{e^{-t/\tau} \cdot k \cdot e^{a(m-M_c)}}{(t+c)^{1+\omega} \cdot \left(r^2 + d \cdot e^{\gamma(m-M_c)}\right)^{1+\rho}},$$
(6)

where r^2 is the squared distance between events and $k, a, c, \omega, \tau, d, \gamma, \rho$ are the learnable parameters along with the constant background rate μ . This triggering kernel is derived from statistical distributions found through decades of observational studies (Utsu & Seki, 1955; Utsu, 1970; Utsu et al., 1995) and several of the learnable parameters have been linked to physical properties of the earthquake rupture process (Utsu et al., 1995; Ide, 2013).

3 EARTHQUAKENPP DATASETS

The EarthquakeNPP datasets encompass earthquake records, including timestamps, geographical coordinates, and magnitudes, documented within California from 1971 to 2021. California, with its dense network and high seismic hazard, has been extensively studied, demonstrating the utility of forecasting algorithms (Gerstenberger et al., 2004; Field, 2007; Field et al., 2021). It encompasses the San Andreas fault plate boundary system (Zoback et al., 1987) and includes modern high-resolution catalogs with numerous small magnitude earthquakes, offering potential for new, more expressive models.

Notebooks to access and preprocess these public datasets along with the associated bench marking experiment are publicly accessible at https://anonymous.4open.science/r/
 EarthquakeNPP-2D51, accompanied by more detailed documentation for each dataset. A summary of how earthquake datasets are generated, along with the associated challenges of using earthquake catalog data can be found in Appendix A. Table 2 provides a short summary of each EarthquakeNPP dataset.

BENCHMARKING EXPERIMENT

We now use EarthquakeNPP to benchmark three spatio-temporal NPPs with prior positive claims on earthquake forecasting.

Neural Spatio-Temporal Point Process (NSTPP) (Chen et al., 2021): a pdf based NPP that
 parameterizes the spatial pdf with continuous-time normalizing flows (CNFs). We use their Attentive
 CNF model for its computational efficiency and overall performance versus their other model Jump
 CNF (Chen et al., 2021).



270Table 2: Summary of EarthquakeNPP datasets, including: region, dataset development, magnitude271threshold (M_c), number of training (combined with validation) events, and number of testing events.272The chronological partitioning of training, validation, and testing periods is also detailed. An auxiliary273(burn-in) period begins from the **Start** date, followed by the respective starts of the training, validation,274and testing periods. All dates are given as 00:00 UTC on January 1st, unless noted (* refers to 00:00275UTC on January 17th). Finally, we give our purpose for including each dataset.

	ComCat	SCEDC	White	QTM
Region	Whole of California	Southern California	San Jacinto Fault-Zone	QTM_SanJac: San Jacinto Fault-Zone,
				QTM_SaltonSea: Salton Sea
Development	The U.S. Geological Survey (USGS) National Earthquake Information Center (NEIC) monitors global earthquakes (Mw 4.5 or larger) and provides complete seismic monitoring of the United States for all significant earthquakes (> Mw 3.0 or felt). Its contributing seismic networks have produced the Advanced National Seismic System (ANSS) Comprehen- sive Catalog of Earthquake Events and Products.	The Southern California Seismic Network (SCSN) has developed and main- tained the standard earth- quake catalog for Southern California (Hutton et al., 2010) since the Caltech Seismological Laboratory began routine operations in 1932. Significant net- work improvements since the 1970s and 1980s re- duced the catalog com- pleteness from Mw 3.25 to Mw 1.8.	White et al. (2019) cre- ated an enhanced cata- log focusing on the San Jacinto fault region, us- ing a dense seismic net- work in Southern Califor- nia. This denser network, combined with automated phase picking (STA/LTA), ensures a 99% detec- tion rate for earthquakes greater than Mw 0.6 in a specific subregion (White et al., 2019).	Using data collected by the SCSN, Ross et al. (2019) generated a denser catalog by reanalyzing the same waveform data with a template matching procedure that looks for cross-correlations with the wavetrains of previously detected events.
M_c	Mw 2.5	SCEDC_20: Mw 2.0, SCEDC_25: Mw 2.5, SCEDC_30: Mw 3.0	Mw 0.6	Mw 1.0
# Train/Test Events	79,037 / 23,059	SCEDC_20: 128,265 / 14,351, SCEDC_25: 43,221 / 5,466, SCEDC_30: 12,426 / 2,065	38,556 / 26,914	QTM_SanJac: 18,664 / 4,837, QTM_SanJac: 44,042 / 4,393
Start-Train- Val-Test-End	1971-1981-1998-2007- 2020*	1981-1985-2005-2014- 2020	2008-2009-2014-2016- 2018	2008-2009-2014-2016- 2018
Purpose	Example of data currently in use for operational forecasting (USGS utilizes ComCat in aftershock fore- casts they release to the pub- lic.)	Three magnitude thresh- olds (Mw 2.0, 2.5, 3.0) ex- plore the effect of trunca- tion on forecasting model performance.	To explore if newly detected low magnitude earthquakes contain additional predictive information.	To explore if newly de- tected low magnitude earthquakes contain ad- ditional predictive infor- mation (with different detection methodology).

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Deep Spatio-Temporal Point Process (Deep-STPP) (Zhou et al., 2022): a conditional intensity
 function based NPP that constructs a non parametric space-time intensity function governed by a
 deep latent process. The intensity function enjoys a closed form integration, avoiding the need for
 numerical approximation.

Automatic Integration for Spatiotemporal Neural Point Processes (AutoSTPP) (Zhou & Yu, 2024): a conditional intensity function based NPP which jointly models the 3D space-time integral of the intensity along with its derivative (the intensity function) using a dual network approach.

The benchmark is against the **ETAS** model defined in section 2.2, as well as a homogeneous **Poisson** process. The Poisson model is fit to events in the auxiliary, training and validation windows to provide a baseline score against which to compare all four other models.

Validation is typically not part of the estimation procedure for ETAS, so it is fit using the combined
 training and validation windows. NPPs follow the standard training/validation/testing procedure of
 machine learning. When possible, a model's likelihood for training, validation, and testing can depend
 on events occurring before the splits through memory in its history. The exception is NSTPP, lacking



Figure 2: Test temporal log-likelihood scores for all the spatio-temporal point process models on each of the EarthquakeNPP datasets. Error bars of the mean and standard deviation are constructed for the NPPs using three repeat runs.

a direct dependency on prior events. Nonetheless, its likelihood is evaluated on the same events
as the other models. The definition of the ETAS model (equation 4) specifies how the magnitudes
of earthquakes in the history contribute towards the intensity function. This earthquake magnitude
dependence is not implemented in any of the NPPs we benchmark, since it requires modeling choices
beyond the scope of this work.

354 Figures 2 and 3 present the temporal and spatial log-likelihood scores of all models on the Earthquak-355 eNPP datasets. The ETAS model consistently achieves the highest temporal and spatial log-likelihood 356 across all datasets, though some NPP models demonstrate comparable temporal performance on 357 the ComCat, QTM_SaltonSea, QTM_SanJac, and White catalogs. Among the NPP models, 358 Deep-STPP generally exhibits the best temporal log-likelihood, likely due to its formulation, which accounts for the influence of unobserved events—a phenomenon that varies temporally in earthquake 359 data (see Section A.2). In contrast, AutoSTPP achieves the highest spatial log-likelihood, attributed 360 to its ability to capture anisotropic Hawkes kernels (see Figure 2 of Zhou & Yu (2024)), which are 361 often observed in earthquake data (Page & van der Elst, 2022). 362

The improved relative temporal performance of all NPPs compared to ETAS, particularly when the magnitude threshold is lowered from 3.0 to 2.0 in the SCEDC dataset, indicates that low magnitude 364 earthquakes provide valuable predictive information for NPPs. This is further suggested by the 365 comparable performance of NPPs to ETAS on low-magnitude catalogs such as QTM_SaltonSea, 366 QTM_SanJac, and White. The stronger temporal performance of NPPs on datasets such as 367 ComCat, QTM_SaltonSea, QTM_SanJac, and White may also reflect their ability to model 368 more complex physical processes, such as earthquake swarms (Llenos & van der Elst, 2019) or 369 tectonic activity near the Mendocino Triple Junction (Hellweg et al., 2024). Additional datasets and 370 results are included in Appendix B.

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5 CSEP CONSISTENCY TESTS

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EarthquakeNPP also supports the earthquake forecast evaluation protocol developed by the Collabora tory for the Study of Earthquake Predictability (CSEP). In this procedure a model generates 24-hour
 forecasts through 10,000 repeat simulations of earthquake sequences at the beginning of every day in
 the testing period. This procedure mimics how earthquake forecasts are generated in an operational



Figure 3: Test spatial log-likelihood scores for all the spatio-temporal point process models on each of the EarthquakeNPP datasets. Error bars of the mean and standard deviation are constructed for the NPPs using three repeat runs.

setting (van der Elst et al., 2022). Models can then be evaluated by comparing the observed sequence with the distribution over model simulations. Three test statistics target the temporal, spatial and magnitude components of the forecasts, where a test is failed if the observed statistic falls within a pre-defined rejection region (Figure 4). We demonstrate this procedure for the ETAS model and report performance scores as a benchmark for future implementations of NPPs. A case study using the 2019 M7.1 Ridgecrest earthquake can be for found in Appendix F.

5.1 NUMBER (TEMPORAL) TEST

The number test evaluates the temporal component of the forecast by checking the consistency of the forecasted number of events, N with those observed in the forecast horizon, N_{obs} . Upper and lower quantiles are estimated using the empirical cumulative distribution from the repeat simulations, F_N ,

$$\delta_1 = \mathbb{P}(N \ge N_{\text{obs}}) = 1 - F_N(N_{\text{obs}} - 1) \tag{7}$$

$$\delta_2 = \mathbb{P}(N \le N_{\text{obs}}) = F_N(N_{\text{obs}}). \tag{8}$$

5.2 SPATIAL TEST

To evaluate the spatial component of the forecast, a test statistic aggregates the forecasted rates of earthquakes over a regular grid,

$$S = \left[\sum_{i=1}^{N} \log \hat{\lambda}(k_i)\right] N^{-1},\tag{9}$$

where $\hat{\lambda}(k_i)$ is the approximate rate in the cell k where the i^{th} event is located. Upper and lower quantiles are estimated by comparing the observed statistic

$$S_{\text{obs}} = \left[\sum_{i=1}^{N_{\text{obs}}} \log \hat{\lambda}(k_i)\right] N_{\text{obs}}^{-1},\tag{10}$$

with the empirical cumulative distribution of S using the repeat simulations, F_S

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$$\gamma_s = \mathbb{P}(S \le S_{\text{obs}}) = F_S(S_{\text{obs}}). \tag{11}$$



Figure 4: CSEP consistency tests on the ETAS model for the first day (01/01/2014) of the testing period in the SCEDC catalog. A total of 10,000 simulations are generated to compute empirical distributions of the test statistics for each of the three consistency tests: (a) Number test, (b) Spatial test, and (c) Magnitude test. The test fails if the observed statistic falls within the rejection region (red), defined by the 0.05 and 0.95 quantiles of the distribution.

The grid is constructed from $\{0.1^\circ, 0.05^\circ, 0.01^\circ\}$ squares for ComCat, SCEDC and $\{QTM_Salton_Sea, QTM_SanJac, White\}$ respectively.

5.3 MAGNITUDE TEST

To evaluate the earthquake magnitude component of the forecast, a test statistic compares the histogram of a forecast's magnitudes $\Lambda^{(m)}$, against the mean histogram over all forecasts $\bar{\Lambda}^{(m)}$,

$$D = \sum_{k} \left(\log \left[\bar{\Lambda}^{(m)}(k) + 1 \right] - \log \left[\Lambda^{(m)}(k) + 1 \right] \right)^2, \tag{12}$$

where $\Lambda^{(m)}(k)$ and $\bar{\Lambda}^{(m)}(k)$ are the counts in the k^{th} bin of the forecast and mean histograms, normalised to have the same total counts as the observed catalog. Upper and lower quantiles are estimated by comparing the observed statistic

$$D_{\text{obs}} = \sum_{k} \left(\log \left[\bar{\Lambda}^{(m)}(k) + 1 \right] - \log \left[\Lambda^{(m)}_{\text{obs}}(k) + 1 \right] \right)^2, \tag{13}$$

with the empirical distribution of D using the repeat simulations, F_D

$$\gamma_m = \mathbb{P}(D \le D_{\text{obs}}) = F_D(D_{\text{obs}}). \tag{14}$$

Histogram bins of size $\delta_m = 0.1$ are used across all datasets.

5.4 EVALUATING MULTIPLE FORECASTING PERIODS

Savran et al. (2020) describe how to assess a model's performance across the multiple days in the testing period. By construction, quantile scores over multiple periods should be uniformly distributed if the model is the data generator (Gneiting & Katzfuss, 2014). Therefore comparing quantile scores against standard uniform quantiles (y = x), highlights discrepancies between the observed data and the forecast. Additional statements can be made about over-prediction or under-prediction of each test statistic (quantile curves above/bellow y=x respectively). The Kolmogorov-Smirnov (KS) statistic then quantifies the degree of difference to the uniform distribution for each of the tests.

Further documentation of how to perform the CSEP evaluation procedure can be found on the platform, where we demonstrate the procedure for the ETAS model. Table 3 reports the benchmark performance scores taken from the quantile plots in Appendix D. The performance of ETAS is higher for the more typical higher magnitude catalogs such as ComCat and SCEDC, whereas it performs worse at the lower magnitude catalogs of QTM_SanJac, QTM_SaltonSea and White. Spatial prediction is consistently the best performing component of the ETAS forecast, whereas earthquake numbers are overpredicted by the model and earthquake magnitudes are generally not well predicted (Figure 9). All results indicate significant room for improvement beyond the predictive performance of the ETAS model.

Table 3: CSEP consistency tests evaluate the calibration of all daily ETAS forecasts on EarthquakeNPP datasets. A test is performed at the $\alpha = 0.05$ significance level on each day in the testing period. The pass rate indicates the success of ETAS across all testing days. By construction quantile scores of the tests should be uniformly distributed if the model is the data generator. The KS-Statistic reports the difference of the quantile distribution to uniform, taken from the quantile plots in Appendix D.

Dataset	Num	ber Test	Spat	Spatial Test Magnitude Test		tude Test
	Pass Rate	KS-Statistic	Pass Rate	KS-Statistic	Pass Rate	KS-Statistic
ComCat	62.3%	0.392	85.3%	0.128	75.3%	0.318
SCEDC	74.4%	0.161	87.5%	0.123	80.5%	0.153
QTM_SanJac	59.2%	0.461	96.7%	0.145	66.2%	0.406
QTM_SaltonSea	54.2%	0.441	82.1%	0.216	79.0%	0.311
White	17.1%	0.750	98.0%	0.373	25.0%	0.741

6 DISCUSSION AND CONCLUSION

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We introduce the EarthquakeNPP datasets to facilitate the benchmarking of NPPs against a community-endorsed ETAS model for earthquake forecasting. These datasets cover various regions of California, representing typical forecasting zones and the data commonly available to forecast issuers. Several datasets use modern methods of detection, which enables the inclusion of much smaller magnitude earthquakes.

510 In a benchmarking experiment, we compared three NPP models against ETAS and a baseline Poisson 511 process. None of the NPP models outperformed ETAS, indicating that current NPP implementations 512 are not yet suitable for operational earthquake forecasting. ETAS explicitly defines how larger 513 earthquake magnitudes increase the likelihood of future earthquakes in both time and space, using an 514 empirical relationship derived from seminal observational studies (Utsu & Seki, 1955; Utsu, 1970). 515 This use of magnitude information is shared across all competitive short-term earthquake forecasting models currently used operationally (Mizrahi et al., 2024a) or tested by CSEP (Taroni et al., 2018). 516 The lack of a direct dependence on magnitudes in the current NPP implementations likely explains 517 their relative under-performance compared to ETAS. Future implementations should exploit this 518 additional feature for improved temporal and spatial performance. Encouragingly, the comparable 519 temporal performance to ETAS without this additional feature suggests that incorporating magnitude 520 dependence would enhance NPP performance beyond that of ETAS. 521

EarthquakeNPP supports the earthquake forecast evaluation procedure developed by the Collaboratory 522 for the Study of Earthquake Predictability (CSEP). The procedure replicates how earthquakes forecasts 523 are generated in an operational setting, requiring models to simulate many repeat event sequences over 524 a day-long forecast horizon. Benchmark performance for the ETAS model enables future comparison 525 of NPPs that are implemented for this procedure and enables their promotion to the fully prospective 526 CSEP experiments. Notably, this procedure allows the evaluation of generative NPP models without 527 explicit likelihoods (Yuan et al., 2023; Li et al.), by assessing their performance over the full trajectory 528 of future events. Probabilistic seismic hazard analysis (PSHA) requires long-term prediction beyond 529 the next-event (Ebrahimian et al., 2014; Gerstenberger et al., 2014), therefore this approach also 530 offers stakeholders a more comprehensive understanding of earthquake hazard than metrics focused 531 on predicting the next event (e.g. RMSE). The procedure also follows the recommendation by Shchur et al. (2021) to move away from next-event point prediction for NPPs. 532

The EarthquakeNPP datasets, available at https://anonymous.4open.science/r/
 EarthquakeNPP-2D51, provide a platform for future NPP developments to be benchmarked against these initial results. The platform is under ongoing development and in the future will see the direct comparison of emerging and other existing models models developed within the seismology community, as well as an expansion of datasets included to other seismically active global regions.
 Successful NPP models on these datasets, for both log-likelihood and CSEP metrics, will be directly impactful to stakeholders in seismology, ultimately enabling their integration into operational earthquake forecasting by government agencies.

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A EARTHQUAKE CATALOG DATA

A.1 EARTHQUAKE CATALOG GENERATION



Figure 5: Generating an earthquake catalog involves several key steps: seismic phase picking, magnitude estimation, and the association and location of seismic sources. This process transforms raw waveform data recorded at seismic stations to locations, times, and magnitudes of earthquakes.

Data missingness, referred to in seismology as catalog (in)completeness, is the primary challenge
 faced with earthquake catalogs. It is an important and unavoidable feature, and is a result of how
 earthquakes are detected and characterised. Below, we briefly overview the process of generating an
 earthquake catalog to illustrate the data quality issues. In the subsequent section, we review catalog
 incompleteness and its potential impact on the performance and evaluation of forecast models.

Seismometers and Seismic Networks. A seismometer is an instrument that detects and records the vibrations caused by seismic waves (Stein & Wysession, 2009; Shearer, 2019). It consists of a sensor to detect ground motion and a recording system to log three-dimensional ground motion over time, typically vertical and horizontal velocities. Seismic networks, comprising multiple seismometers, monitor seismic activity at regional, national or global scales (see, e.g., (Woessner et al., 2010) and references therein). High-density networks with modern, sensitive equipment provide more detailed and accurate data, enhancing the ability to detect and analyse smaller and more distant earthquakes.

From Waveforms to Phase Picking. The process of converting raw continuous seismic waveforms into useful earthquake data begins with phase picking, which identifies the arrival times of the primary (P) and secondary (S) waves of an earthquake. Historically, this was done manually, but now automated algorithms, such as the STA/LTA algorithm, detect wave arrivals by analyzing signal amplitude changes (Allen, 1982). Recent algorithms, such as machine learning classifiers (e.g. Zhu & Beroza, 2019; Lapins et al., 2021) and template-matching (e.g. Ross et al., 2019), can process much higher volumes of data efficiently and are often able to detect events of much smaller magnitudes.

Earthquake Association and Location After phase picking, the next step is to associate phases from different seismometers with the same earthquake. Simple algorithms require at least four

918 phase arrivals to be detected on different stations within a short time interval to declare an event. 919 Once phases are associated, location estimation determines the earthquake's hypocenter and origin 920 time by minimizing travel-time residuals using linearized or global inversion algorithms (Thurber, 921 1985; Lomax et al., 2000). Given the potential for misidentified or mis-associated phase arrivals 922 due to low signal-to-noise of small events or the near-simultaneous occurrence during very active aftershock sequences, an automated system typically first picks arrival times and determines a 923 preliminary location, which is subsequently reviewed by a seismologist (e.g. Woessner et al., 2010, 924 and references therein). Locations are typically reported as the geographical coordinates and depths 925 where earthquakes first nucleated (hypocenters), although some catalogs report the centroid location, 926 a central measure of the extended earthquake rupture. 927

Earthquake Magnitude Calculation The magnitude of an earthquake quantifies the energy released at the source and was originally defined in the seminal paper by Richter (1935). The original definition, now referred to as the local magnitude (ML), is calculated from the logarithm of the amplitude of waves recorded by seismometers. This scale, however, "saturates" at higher magnitudes, meaning it underestimates magnitudes for various reasons. This led to introduction of the moment magnitude scale (Mw) (Hanks & Kanamori, 1979), which computes the magnitude based on the estimated seismic moment M_0 , which can be related to the physical rupture process via

$$M_0 = \text{rigidity} \times \text{rupture area} \times \text{slip}, \tag{15}$$

where rigidity is a mechanical property of the rock along the fault, rupture area is the area of the fault that slipped, and slip is the distance the fault moved. Mw is determined seismologically via a spectral fitting process to the earthquake waveforms. In practice, it can be challenging to use a single magnitude scale for a broad range of magnitudes, therefore a range of scales may be present within a single catalog, and approximate magnitude conversion equations may be used to homogenize the scales (e.g. Herrmann & Marzocchi, 2021, and references therein).

943 944 A.2 EARTHQUAKE CATALOG COMPLETENESS

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All of the EarthquakeNPP datasets are made publicly available by their respective data centers in raw format. However, constructing a suitable retrospective forecasting experiment from this raw data requires appropriate pre-processing. This typically involves truncating the dataset above a magnitude threshold $M_{\rm cut}$ and within a target spatial region to address incomplete data, known as catalog completeness M_c (e.g., Mignan et al., 2011; Mignan & Woessner, 2012).

950 There are several reasons why an earthquake may not be detected by a seismic network. Small events 951 may be indistinguishable from noise at a single station, or insufficiently corroborated across multiple 952 stations. Another significant cause of missing events occurs during the aftershock sequence of large 953 earthquakes, when the seismicty rate is high (Kagan & Knopoff, 1987; Hainzl, 2022). Human or 954 algorithmic detection abilities are hampered when numerous events occur in quick succession, e.g. 955 when phase arrivals of different events overlap at different stations or the amplitudes of small events are swamped by those of large events. Since catalog incompleteness increases for lower magnitude 956 events, typically the task is to find the value M_c above which there is approximately 100% detection 957 probability. Choosing a truncation threshold $M_{\rm cut}$ that is too high removes usable data. Where 958 NPPs have demonstrated an ability to perform well with incomplete data (Stockman et al., 2023), 959 typically a threshold below the completeness biases classical models such as ETAS (Seif et al., 2017). 960 Seismologists often investigate the biases of different magnitude thresholds by performing repeat 961 forecasting experiments for different thresholds (e.g. Mancini et al., 2022; Stockman et al., 2023), 962 which we also facilitate in our datasets. 963

Typically M_c is determined by comparing the raw earthquake catalog to the Gutenberg-Richter law (Gutenberg & Richter, 1936), which states that the distribution of earthquake magnitudes follows an exponential probability density function

$$f_{GR}(m) = \beta e^{\beta(m-M_c)} \quad : m \ge M_c.$$
(16)

969 where β is a rate parameter related to the b-value by $\beta = b \log 10$. Histogram-based approaches, 970 such as the simple Maximum Curvature method (Wiemer & Wyss, 2000) as well as many others (e.g. 971 Herrmann & Marzocchi, 2021, and references therein), identify the magnitude at which the observed 972 catalog deviates from this law, indicating incompleteness (See Figure 6b).

972 In practice, catalog completeness varies in both time and space $M_c(t, \mathbf{x})$ (e.g. Schorlemmer & 973 Woessner, 2008). During aftershock sequences, $M_c(t)$ can be very high (e.g., Agnew, 2015; Hainzl, 974 2016b) (See Figure 6a). Thresholding at the maximum value might remove too much data. Instead, 975 modelers either omit particularly incomplete periods during training and testing (Kagan, 1991; Hainzl 976 et al., 2008), model the incompleteness itself (Helmstetter et al., 2006; Werner et al., 2011; Omi et al., 2014; Hainzl, 2016a;b; Mizrahi et al., 2021; Hainzl, 2022), or accept known biases from 977 disregarding this issue (Sornette & Werner, 2005). Spatially, catalogs are less complete farther from 978 the seismic network (Mignan et al., 2011), so the spatial region can be constrained to remove outer, 979 more incomplete areas (See Figure 6c). 980



Figure 6: a) the June 10, 2016 Mw5.2 Borrego Springs earthquake and aftershocks, which occurred 1008 on the San Jacinto fault zone and is recorded in the WHITE catalog. An estimate of the magnitude of 1009 completeness $M_c(t)$ over time using the Maximum Curvature method reveals more incompleteness 1010 immediately following the large earthquake. b) magnitude-frequency histograms reveal that truncating 1011 the raw WHITE catalog to inside the target region decreases M_c . Each histogram is fit to the Gutenberg-1012 Richter (GR) law and an estimate of M_c for each catalog occurs where the histogram deviates from 1013 the (GR) line. c) An estimate of M_c for gridded regions of the San Jacinto fault zone, using the raw 1014 WHITE catalog. 1015

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1017 B ADDITIONAL DATASETS

Beyond the official EarthquakeNPP datasets, we include 3 further datasets that either provide additional scientific insight or continuity from previous benchmarking works.

Synthetic ETAS Catalogs. We simulate a synthetic catalog using the ETAS model with parameters estimated from ComCat, at M_c 2.5, within the same California region. A second catalog emulates the time-varying data-missingness present in observational catalogs by removing events using the time-dependent formula from Page et al. (2016),

$$M_c(M,t) = M/2 - 0.25 - \log_{10}(t), \tag{17}$$

1026 Table 4: Summary of additional datasets, including: magnitude threshold (M_c), number of training 1027 events, and number of testing events. The chronological partitioning of training, validation, and 1028 testing periods is also detailed. An auxiliary (burn-in) period begins from the "Start" date, followed by the respective starts of the training, validation, and testing periods. All dates are given as 00:00 1029 UTC on January 1st, unless noted (* refers to 00:00 UTC on January 17th). 1030

$\mathbf{M}_{\mathbf{c}}$	Start-Train-Val-Test-End	Train Events	Test Events
2.5	1971-1981-1998-2007-2020*	117,550	43,327
2.5	1971-1981-1998-2007-2020*	115,115	42,932
2.5	1990-1992-2007-2011-2020	22,213	15,368
	M _c 2.5 2.5 2.5	Mc Start-Train-Val-Test-End 2.5 1971-1981-1998-2007-2020* 2.5 1971-1981-1998-2007-2020* 2.5 1990-1992-2007-2011-2020	Mc Start-Train-Val-Test-End Train Events 2.5 1971-1981-1998-2007-2020* 117,550 2.5 1971-1981-1998-2007-2020* 115,115 2.5 1990-1992-2007-2011-2020 22,213

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where M is the mainshock magnitude. Events below this threshold are removed using mainshocks of 1043 Mw 5.2 and above. The inclusion of these datasets allows us to test whether NPPs are inhibited by 1044 data missingness to the same extent that ETAS is.

1045 **Deprecated Catalog of Japan.** To provide continuity from the previous benchmarking for NPPs 1046 on earthquakes, we also provide results on the Japanese dataset from Chen et al. (2021), however 1047 with a chronological train-test split and without removing any supposed outlier events. To reflect our 1048 recommendation not to use this dataset in any future benchmarking following the dataset completeness 1049 issues mentioned above, we name this dataset Japan Deprecated.

1050 Figures 7 and 8 report the temporal and spatial log-likelihood scores of all the benchmarked models 1051 on additional datasets. On synthetic data generated by the ETAS model the performance of NPPs 1052 mirrors the results on the observational data (Figures 2 and 3). The performance of NPPs is more 1053 comparable to ETAS in terms of temporal log-likelihood however they cannot capture the distribu-1054 tion of earthquake locations. Change in temporal performance of models between the ETAS and 1055 ETAS_incomplete datasets reveal each model's robustness to the missing data typically present 1056 in earthquake catalogs (See section A.2). Auto-STPP and ETAS reduce in performance upon the removal earthquakes during aftershock sequences, whereas DeepSTPP and NSTPP maintain the same 1057 performance indicating a robustness to the data missingness. 1058

1059 On the Japan Deprecated dataset, whilst ETAS remains the best performing model for spatial prediction, for temporal prediction it performs comparably to NSTPP and is even marginally out-1061 performed by DeepSTPP. This performance can be attributed to the data completeness issues of the 1062 Japan Deprecated dataset (see section 1.1), where the test period is missing all earthquakes bellow magnitude 4.0. 1063









Figure 8: Test spatial log-likelihood scores for all the spatio-temporal point process models on each of the additional datasets. Error bars of the mean and standard deviation are constructed for the NPPs using three repeat runs.

COMPUTATIONAL EFFICIENCY С

1098 C.1 TRAINING 1099

1100 Table 5 reports the training times for each model across all datasets. We ran all the NPP models using 1101 a HPC node with Nvidia Ampere GPU with 4x Nvidia A100 40GB SXM "Ampere" GPUs and AMD 1102 EPYC 7543P 32-Core Processor "Milan" CPU using torch==1.12.0 and cuda==11.3. 1103

Dataset	# Training Events	ETAS	Deep-STPP	AutoSTPP	NSTPP	Poisson
ComCat	79,037	08:59:04	00:15:35	01:34:09	3 days, 05:10:17	<1 second
QTM_SaltonSea	44,042	07:28:28	00:26:46	01:45:34	2 days, 00:26:45	<1 second
QTM_SanJac	18,664	00:32:40	00:09:31	00:37:03	1 day, 22:06:33	<1 second
SCEDC_20	128,265	13:42:30	00:38:10	02:54:51	3 days, 02:20:40	<1 second
SCEDC_25	43,221	03:09:14	00:09:34	00:56:05	2 days, 16:33:55	<1 second
SCEDC_30	12,426	00:42:25	00:02:44	00:16:01	1 day, 16:39:04	<1 second
White	38,556	03:55:40	00:08:21	01:10:51	2 days, 01:03:57	<1 second
Japan_Depreca	ted 22,213	06:09:08	00:13:45	01:02:07	2 days, 05:32:03	<1 second
ETAS	117,550	00:33:25	00:15:24	01:10:22	3 days, 03:09:17	<1 second
ETAS_incomple	te 115,115	00:35:14	00:15:29	01:09:43	3 days, 11:39:51	<1 second

1122 Table 5: Training times for each model across all datasets, including the number of training events. 1123 Times are formatted as HH:MM:SS, with days included for durations exceeding 24 hours. The Poisson model consistently requires less than 1 second. 1124

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ETAS training scales $\mathcal{O}(n^2)$ with the total number of events, since for every event a contribution to 1126 the intensity function is computed from a summation over all previous events. This scaling, coupled 1127 with the lack of parallelization in the current implementation, results in long training times for larger 1128 datasets. Poorer scaling will likely hinder ETAS if dataset sizes continue to grow in the future 1129 (Stockman et al., 2024). 1130

1131 Encouragingly, both **Deep-STPP** and **AutoSTPP** are significantly faster to train due to GPU acceleration and their use of a sliding window of the most recent k = 20 events. While exact complexity 1132 analyses are not provided in Zhou et al. (2022) or Zhou & Yu (2024), we can infer that Deep-STPP 1133 likely scales as $\mathcal{O}(kn)$ since it benefits from a closed-form expression for the likelihood. AutoSTPP,

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though requiring automatic integration to compute the likelihood, still scales with O(kn) because the additional integration cost does not affect the overall scaling.

NSTPP, on the other hand, incurs significant training costs, rendering it impractical for real-time forecasting. Unlike the sliding window mechanism used in **Deep-STPP** and **AutoSTPP**, **NSTPP** partitions the event sequence into fixed time intervals, leading to sequences that are much longer than the k = 20 events used by the other models (as shown in Figure 11 of Chen et al. (2021)). Furthermore, solving an ODE for each event time adds a significant computational burden, even with the use of their faster attentive CNF architecture.

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C.2 SIMULATION

1145 Real-time earthquake forecasting and CSEP model evaluation require simulating many repeat se-1146 quences (at least 10,000 for adequate distributional coverage) over the forecasting horizon. While ETAS training scales as $\mathcal{O}(n^2)$ with the number of training events, its simulation scales more ef-1147 ficiently at $\mathcal{O}(n \log n)$. This improved scaling is due to its equivalent formulation as a Hawkes 1148 branching process (see Section 2.2). Both Deep-STPP and AutoSTPP are also based on Hawkes 1149 processes, which theoretically allows for fast simulation. However, as these models currently only 1150 have an intensity function implementation, simulating events would require a slower thinning pro-1151 cedure (Ogata, 1981), limiting their simulation efficiency. In contrast, NSTPP benefits from fast 1152 simulation, owing to its design using continuous-time normalizing flows. Events can be generated by 1153 passing samples from a base distribution through learned transformations, resulting in a much faster 1154 simulation process.

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D CSEP CONSISTENCY TESTS

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Figure 9: Quantile-quantile plots showing the calibration of all daily ETAS forecasts on a) ComCat,b) SCEDC, c) QTM_San_Jac, d) QTM_Salton_Sea, e) White. By construction quantile scores over multiple periods should be uniformly distributed if the model is the data generator. Comparing quantile scores against standard uniform quantiles (y = x), highlights discrepancies between the observed data and the forecast. Pass rates of each test are indicated in the legend. The Kolmogorov-Smirnov statistic, quantifies the degree of difference to the uniform distribution.



Figure 10: Times and magnitudes of events in the ComCat dataset (with key events labeled). The size of the points are plotted on a log scale corresponding to Mw. Auxiliary, training, validation and testing periods are indicated by colour and a further cumulative count of events is indicated in red.



Figure 11: Locations of events in the ComCat dataset, labeled by their partition into auxiliary, training, validation and testing periods.



Figure 12: Times and magnitudes of events in the SCEDC dataset (with key events labeled). The size of the points are plotted on a log scale corresponding to Mw. Auxiliary, training, validation and testing periods are indicated by colour and a further cumulative count of events is indicated in red.



Figure 13: Locations of events in the SCEDC dataset, labeled by their partition into auxiliary, training, validation and testing periods.



Figure 14: Times and magnitudes of events in the White dataset (with key events labeled). The size of the points are plotted on a log scale corresponding to Mw. Auxiliary, training, validation and testing periods are indicated by colour and a further cumulative count of events is indicated in red.



Figure 15: Locations of events in the White dataset, labeled by their partition into auxiliary, training, validation and testing periods.



Figure 16: Times and magnitudes of events in the QTM_SanJac dataset. The size of the points are plotted on a log scale corresponding to Mw. Auxiliary, training, validation and testing periods are indicated by colour and a further cumulative count of events is indicated in red.





Figure 17: Times and magnitudes of events in the QTM_SaltonSea dataset. The size of the points are plotted on a log scale corresponding to Mw. Auxiliary, training, validation and testing periods are indicated by colour and a further cumulative count of events is indicated in red.



Figure 18: Locations of events in the QTM_SanJac and QTM_SaltonSea datasets, labeled by their partition into auxiliary, training, validation and testing periods.

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F 2019 M7.1 RIDGECREST EARTHQUAKE CASE STUDY

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The 2019 Ridgecrest earthquake sequence (Figure 19) was the most powerful seismic event to strike
Southern California in the past 20 years. Centered near the town of Ridgecrest and the Naval Air
Weapons Station China Lake, the sequence began with a magnitude 6.4 foreshock on July 4, 2019, at
17:33:49 UTC, followed by a more powerful magnitude 7.1 mainshock on July 6, 2019, at 03:19:53
UTC, both along the Eastern California Shear Zone. The earthquakes caused widespread surface
rupture, with displacements along multiple faults, and triggered tens of thousands of aftershocks over
the following months.

The impacts of the sequence were substantial. In Ridgecrest and surrounding areas, the shaking damaged homes, businesses, and infrastructure, including roads, water lines, and electrical systems. Fires broke out due to ruptured gas lines, exacerbating the destruction. The mainshock caused over \$1 billion in damages, including significant damage to the China Lake Naval facility, which was temporarily evacuated and declared "not mission capable." Despite the severity of the shaking, no fatalities occurred, largely due to the remote location and earthquake-resistant construction in the region.

Using the CSEP evaluation procedure (Section 5), we isolate the performance of a model during the sequence to identify its strengths and weaknesses. Here, we apply this analysis to the ETAS model, illustrating how similar evaluations can be conducted for future implementations of NPPs or other machine learning-based models.

Figure 20 presents the results of the Number Test over the initial days of the sequence. ETAS forecasts consistently underestimate the number of aftershocks during the most seismically active phase of the sequence. It is only 4 days after the M7.1 Ridgecrest mainshock, that ETAS begins to provide accurate earthquake rate forecasts. Figure 21a shows the spatial forecast for the day after the M7.1





Figure 19: The 2019 Ridgecrest earthquake sequence began with the M6.4 Searles Valley foreshock on July 4, 2019, at 17:33:49 UTC, followed by the M7.1 Ridgecrest mainshock on July 6, 2019, at 03:19:53 UTC. (a) The times and magnitudes of events in the sequence. (b) Events in the sequence are plotted on a map of modeled faults in California.



Figure 20: Forecasted earthquake number distribution using the ETAS model during the first 10 days
of the Ridgecrest earthquake sequence. The number distributions are generated through 10,000 repeat
simulations of earthquake sequences from the beginning of the day. The 95% confidence interval
of the forecasted counts, generated at the start of each day, is compared to the observed number of
events recorded by the end of the day.



Figure 21: (a) The forecasted rates of earthquakes on July 7 (the day after the M7.1 Ridgecrest earthquake) using the ETAS model. Rates are calculated through 10,000 repeat simulations of earthquake sequences from the beginning of the day, which are aggregated to estimate a rate per spatial grid cell. In red are the observed earthquakes that occurred that day. (b) The results of the Spatial Test for July 7. Since the observed statistic is well outside the forecast distribution, the test is failed.



¹⁶²⁰ G ERROR DISTRIBUTIONS & NEXT-EVENT METRICS

Figure 22: The distribution of errors $(Y_{obs} - Y_{pred})$ for the Normal(0, 1), Exponential(1), and Pareto(2.5) distributions. Maximum likelihood estimation is used to fit Normal and Laplace distributions to each error histogram. Normal errors (Normal × Normal) are best approximated by the Root Mean Square Error (RMSE), while Laplacian errors (Exponential × Exponential) are best approximated by the Mean Absolute Error (MAE). However, neither RMSE nor MAE effectively capture the errors for the heavy-tailed Pareto distribution.