### **000 001 002 003** IMPROVED RISK BOUNDS WITH UNBOUNDED LOSSES FOR TRANSDUCTIVE LEARNING

Anonymous authors

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## ABSTRACT

In the transductive learning setting, we are provided with a labeled training set and an unlabeled test set, with the objective of predicting the labels of the test points. This framework differs from the standard problem of fitting an unknown distribution with a training set drawn independently from this distribution. In this paper, we primarily improve the generalization bounds in transductive learning. Specifically, we develop two novel concentration inequalities for the suprema of empirical processes sampled without replacement for unbounded functions, marking the first discussion of the generalization performance of unbounded functions in the context of sampling without replacement. We further provide two valuable applications of our new inequalities: on one hand, we firstly derive fast excess risk bounds for empirical risk minimization in transductive learning under unbounded losses. On the other hand, we establish high-probability bounds on the generalization error for graph neural networks when using stochastic gradient descent which improve the current state-of-the-art results.

- 1 INTRODUCTION
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**029 030 031 032 033 034 035 036 037 038** In the field of machine learning research, the analysis of stochastic behavior based on empirical processes is an essential component of learning theory, particularly in understanding and enhancing algorithm performance. The supremum of empirical processes plays a crucial role in various application scenarios, such as empirical process theory, Rademacher complexity theory, Vapnik–Chervonenkis theory, etc. In recent years, concentration inequalities for traditional suprema of empirical processes are fully established fields and have been well studied in the literature such as [\[36,](#page-10-0) [4,](#page-9-0) [5,](#page-9-1) [1,](#page-9-2) [24,](#page-10-1) [39,](#page-11-0) [12,](#page-9-3) [29\]](#page-10-2). All these inequalities based on the assumption of independent and identically distributed random variables. However, in many practical contexts, the i.i.d. assumption does not hold, such as when training and testing data are drawn from different distributions or when there is temporal dependence among data points. Such scenarios are prevalent in fields like visual recognition and computational biology, necessitating alternatives to Talagrand's inequality.

**039 040 041 042 043 044 045 046 047 048** Another significant context in learning theory is transductive learning which was firstly introduced by [\[40\]](#page-11-1). In transductive learning, the training samples are independently and without replacement drawn from a finite population, as opposed to the classic model of independent and with replacement sampling. In this setting, the learning algorithm not only acquires a labeled training set but also receives a set of unlabeled testing instances, with the goal of accurately predicting the labels of the test points. This configuration naturally arises in numerous applications such as text mining, computational biology, recommendation systems, visual recognition, and malware detection. In these cases, the number of unlabeled samples often far exceeds that of labeled samples, and the cost of labeling the unlabeled samples is high. Consequently, the development of transductive algorithms that leverage unlabeled data to enhance learning performance has increasingly attracted attention.

**049 050 051 052 053** In theoretical analysis of transduction learning, we need to sample without replacement, which leads to big challenge and has not been fully understood yet. [\[13\]](#page-9-4) firstly extended the global Rademacher complexities into transductive learning and established the inequalities without replacement. [\[38\]](#page-10-3) derived two concentration inequalities using Hoeffding's reduction method and the entropy method. Nevertheless, both [\[13\]](#page-9-4) and [\[38\]](#page-10-3) considered only bounded function. In real scenarios, where the maximum value of the function may be large and even unbounded, but the frequency of very large

**054 055 056** values tends to be small. To the best of our knowledge, the analysis in unbounded functions random variables in transductive learning has not been studied yet.

**057 058 059 060 061** In this paper, we focus on sampling without replacement with unbounded functions. We introduce a novel concentration inequality for empirical process upper bounds under the scenario of sampling without replacement, particularly for the case of unbounded functions. This represents the first attempt to discuss generalization performance for unbounded functions under the condition of sampling without replacement.

**062 063 064** In Section [2,](#page-1-0) we provide the definition of the transductive learning set-up, including the basic notations and the discussion of two related transductive learning settings introduced by [\[40\]](#page-11-1). We also introduce the notations of the unbounded random variables used in the following sections.

**065 066 067 068 069 070 071 072** Our new concentration inequalities for the case of unbounded functions are provided in Section [3,](#page-2-0) which are, to the best of our knowledge, the first concentration inequalities for sampling without replacement for classes of unbounded functions. Furthermore, we discuss two significant applications of the new inequalities: firstly, we derive high-probability fast excess risk bounds for unbounded loss in transductive learning based on local uniform convergence in Subsection [4.1;](#page-4-0) secondly, in Subsection [4.2,](#page-5-0) we provide generalization error bounds for Graph Neural Networks (GNNs) with unbounded loss when utilizing Stochastic Gradient Descent (SGD) which is better than the state-of-the-art work [\[37\]](#page-10-4) when  $m = o(N^{2/5})$ . All the proofs in this paper are given in Appendix.

**073** Our contributions are summarized as follows:

- We derive two novel concentration inequalities for suprema of empirical processes when sampling without replacement for classes of sub-Gaussian and sub-exponential functions, which is the first in transductive learning.
- We provide fast excess risk bounds for transductive learning considering Bernstein condition with unbounded losses. To the best of our knowledge, existing results do not provide fast rates in GNNs.
- Applying our inequalities, we obtain the generalization gap of GNNs for node classification task for stochastic optimization algorithm. In more detail, we establish high probability bounds of generalization error and test error under sub-Gaussian and sub-exponential losses. Thanks for considering the variance information, our results are better than [\[37\]](#page-10-4) in some scenarios.

# <span id="page-1-0"></span>2 PRELIMINARIES FOR TRANSDUCTIVE LEARNING

**090 091 092 093 094 095 096 097 098** In transductive learning, the learner is provided with  $m$  labeled training points and  $u$  unlabeled test points. The objective of the learner is to obtain accurate predictions for the test points. Two different settings of transductive learning were given by [\[41\]](#page-11-2). One assumes that both the training and test sets are sampled i.i.d. from a same unknown distribution and the learner is provided with the labeled training and unlabeled test sets. Another assumes that the set  $X_N$  consisting of N arbitrary input points without any other assumptions regrading its underlying source is given. Then we sample  $m \leq N$  objects  $X_m \subseteq X_N$  uniformly without replacement from  $X_N$  which makes the inputs in  $X_m$  dependent. Finally, for each input  $x \in X_m$ , the corresponding output Y from some unknown distribution  $P(Y|X)$ . Thus we obtain all the labels for the set  $X_m$ , we denote the training set as  $S_m = (\mathbf{X}_m, \mathbf{Y}_m)$ . The remaining unlabeled set  $\mathbf{X}_u = \mathbf{X}_N \backslash \mathbf{X}_m$ ,  $u = N - m$  is the test set.

**099 100 101 102 103 104 105 106 107** In this paper we study the second setting, as pointed out by [\[41\]](#page-11-2), any upper generalization bound in the second setting can easily yield a bound for the first setting by just taking expectation. Note that related work [\[10,](#page-9-5) [14,](#page-9-6) [38\]](#page-10-3) considers a special case where the labels are obtained from some unknown but deterministic function  $\phi : \mathcal{X} \mapsto \mathcal{Y}$  so that  $P(\phi(\mathbf{x})|\mathbf{x}) = 1$ . We follow their assumption in this paper. Then the learner is a function model  $f(\mathbf{w})$  w.r.t. the parameters w from some fixed hypothesis parameter space W which may not necessarily containing  $\phi$ . The choice of the learner based on both the labeled training set  $S_m$  and the unlabeled test set  $X_u$ . For brevity, we denote  $\ell(\mathbf{w}; \mathbf{x}) = c(f(\mathbf{w}, \mathbf{x}), \phi(\mathbf{x}))$  w.r.t. the parameters w and the random variable x, where  $c: \mathcal{Y}^2 \mapsto \mathbb{R}_+$ is the cost function to measure the error of predicted label and real label on a point  $X$ . Then we can define the training error and test error of the learner as follows:  $\hat{R}_m(\mathbf{w}) = \frac{1}{m} \sum_{\mathbf{x} \in \mathbf{X}_m} \ell(\mathbf{w}; \mathbf{x})$ ,

<span id="page-2-2"></span>**108 109 110**  $R_u(\mathbf{w}) = \frac{1}{u} \sum_{\mathbf{x} \in \mathbf{X}_u} \ell(\mathbf{w}; \mathbf{x})$ , where hat emphasizes the fact that the training (empirical) error can be computed from the data.

**111 112 113 114 115 116 117 118** For technical reasons that will become clear later, we define the overall error to the union of the training and test sets as  $R_N(\mathbf{w}) = \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{X}_N} \ell(\mathbf{w}; X)$ . The main goal of the learner in transductive setting is to select a proper parameters to minimizing the test error  $R_u(\mathbf{w})$ , which we will denote by  $\mathbf{w}_u^*$ . Since the labels of the test set examples are unknown, we can't compute  $R_u(\mathbf{w})$  and need to estimate it based on the training sample  $X_m$  (and potentially using information from the features  $X_u$ ). A common choice is to replace the test error minimization by empirical risk minimization  $\hat{\mathbf{w}}_m = \arg\min_{\mathbf{w}\in\mathcal{W}}\hat{R}_m(\mathbf{w})$  and to use it as an approximation of  $\mathbf{w}_u^*$ . For  $\mathbf{w}\in\mathcal{W}$  we define the excess risk:

$$
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$$

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$$
\varepsilon_u(\mathbf{w}) = R_u(\mathbf{w}) - \inf_{\mathbf{w}' \in \mathcal{W}} R_u(\mathbf{w}') = R_u(\mathbf{w}) - R_u(\mathbf{w}_u^*).
$$

**121 122 123 124 125** In the following sections, we establish some fundamental notations. We use  $\|\cdot\|_2$  to represent the Euclidean norm of a vector and ∥ · ∥ to denote the spectral norm of a matrix. Throughout this study, we let  $\mathcal{B}(\mathbf{w}'; r) \triangleq {\mathbf{w} : ||\mathbf{w} - \mathbf{w}'||_2 \leq r}$ , representing a ball with center vector w' and radius r. The gradient of the function  $\ell$  with respect to its first argument is denoted as  $\nabla \ell$ . Next, we define the Orlicz norm to describe unbounded random variables.

**126 127 Definition 1** ([\[43\]](#page-11-3)). *For*  $\alpha > 0$ , define the function  $\psi_{\alpha} : \mathbb{R}_+ \to \mathbb{R}_+$  with the formula  $\psi_{\alpha}(x) =$  $\exp(x^{\alpha}) - 1$ *. For a random variable X, define the Orlicz norm* 

$$
||X||_{\psi_{\alpha}} = \inf \{ \lambda > 0 : \mathbb{E} \psi_{\alpha}(|X|/\lambda) \le 1 \}.
$$

**130 131 132** *Furthermore, a random variable*  $X \in \mathbb{R}$  *is sub-Gaussian if there exists*  $K > 0$ *, such that*  $||X||_{\psi_2} \le$ K. A random variable  $X \in \mathbb{R}$  is sub-exponential if there exists  $K > 0$ , such that  $||X||_{\psi_1} \leq K$ . A *random variable*  $X \in \mathbb{R}$  *is sub-Weibull if for*  $\forall \lambda > 0$ *, there exists*  $K > 0$ *, such that*  $||X||_{\psi_\alpha}^{\mathcal{F}} \leq K$ *.* 

**133 134 135 136 137 Remark 1.** Orlicz norm is a classical norm. By choosing an appropriate  $\alpha$ , we can define the tail distribution of random variables to different degrees using the Orlicz norm. This paper mainly discusses sub-Gaussian and sub-exponential distributions for loss functions. We use concentration inequality of the sum for sub-Weibull distribution during some proofs in applications, therefore, we provide this unified definition of unbounded random variables based on the Orlicz norm here.

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## <span id="page-2-0"></span>3 CONCENTRATION INEQUALITIES WITH UNBOUNDED LOSSES

**141 142 143** To gain the generalization error bounds for transductive learning with unbounded losses, we develop the novel concentration inequalities for suprema of empirical processes when sampling without replacement for unbounded functions.

**144 145 146 147 148 149 150** We firstly introduce some necessary notations and settings. Let  $\mathcal{C} = \{c_1, \ldots, c_N\}$  be some finite set. For  $m \leq N$ , let  $\{X_1, \ldots, X_m\}$  and  $\{X'_1, \ldots, X'_m\}$  be sequence of random variables sampled uniformly with and without replacement from C. Let  $\mathcal F$  be a (countable<sup>[1](#page-2-1)</sup> class of functions  $f: \mathcal{C} \to \mathbb{R}$ , such that  $\mathbb{E}[f(X_1)] = 0$  for all  $f \in \mathcal{F}$ . It follows that  $\mathbb{E}[f(X_1')] = 0$  since  $X_1$  and  $X'_1$  are identically distributed. Define the variance  $\sigma^2 = \sup_{f \in \mathcal{F}} \mathbb{V}[f(X_1)]$ . Note that  $\sigma^2 = \sup_{f \in \mathcal{F}} \mathbb{E}[f(X_1)^2] = \sup_{f \in \mathcal{F}} \mathbb{V}[f(X_1')].$  Finally define that the supremum of the empirical process for sampling with and without replacement

$$
Q_m = \sup_{f \in \mathcal{F}} \sum_{i=1}^m f(X_i), \quad Q'_m = \sup_{f \in \mathcal{F}} \sum_{i=1}^m f(X'_i).
$$

**155 156 157 158 159** Concentration inequalities for sampling with replacement  $Q_m$  have undergone extensive investigation, including the exploration of Talagrand-type inequality [\[36\]](#page-10-0) and its variations as presented by [\[5,](#page-9-1) [4\]](#page-9-0). In the case of unbounded functions, certain studies, such as [\[1,](#page-9-2) [12\]](#page-9-3) have established tail bounds through truncation methods and Talagrand-type inequalities for suprema of bounded empirical processes. Nevertheless, as of the current date, no bounds for the suprema of empirical processes

<span id="page-2-1"></span>**<sup>160</sup> 161** <sup>1</sup>Note that all results can be translated to the uncountable classes, for instance, if the empirical process is separable, meaning that  $\mathcal F$  contains a dense countable subset. Details can be referred in page 314 of [\[3\]](#page-9-7) or page 72 of [\[5\]](#page-9-1)

**162 163 164** involving unbounded functions for sampling without replacement  $Q'_m$  have been established in the literature.

**165 166 167** Next, we will introduce the innovative concentration inequalities for the suprema of empirical processes under the condition of sampling without replacement. These new results will be established separately for sub-Gaussian and sub-exponential functions.

Theorem 1. (Concentration inequality when sampling without replacement for classes of sub-**Gaussian functions**) Assume that for all  $c \in C$ ,  $\|\sup_{f \in \mathcal{F}} |f(c)| \|_{\psi_2} < \infty$ , for any  $\epsilon > 0$ , we have the following inequality that

$$
\mathbb{P}\left\{Q'_m - (1+\eta)\mathbb{E}[Q_m] \ge \epsilon\right\} \le 6 \exp\left(-\frac{\epsilon^2}{16(1+\beta)m\sigma^2 + 8C^2 \left\|\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}^2}\right).
$$

We also have that for any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ ,

$$
Q'_m \le (1+\eta)\mathbb{E}[Q_m] + \sqrt{\left(16(1+\beta)m\sigma^2 + 8C^2 \left\|\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}^2\right)\log \frac{6}{\delta}},
$$

where  $\eta$ ,  $\beta$  are some positive constants and C is a positive constants depending on  $\eta$ ,  $\beta$ .

<span id="page-3-0"></span>Theorem 2. (Concentration inequality when sampling without replacement for classes of sub**exponential functions**) Assume for all  $c \in C$ ,  $\|\sup_{f \in \mathcal{F}} |f(c)\|_{\psi_1} < \infty$ , for any  $\epsilon > 0$ , we have the following inequality that

$$
\mathbb{P}\left\{Q'_{m}-(1+\eta)\mathbb{E}[Q_{m}]\geq\epsilon\right\}\leq 2\exp\left(-\frac{\epsilon^{2}}{16(1+\beta)m\sigma^{2}+48C^{2}\left\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f(X_{i})\right\|_{\psi_{1}}^{2}}\right).
$$

We also have that for any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ ,

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$$
Q'_m \le (1+\eta) \mathbb{E}[Q_m] + \sqrt{\left(16(1+\beta)m\sigma^2 + 48C^2 \left\|\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_1}^2\right)\log \frac{2}{\delta}}
$$

,

where  $\eta$ ,  $\beta$  are some positive constants and C is a positive constants depending on  $\eta$ ,  $\beta$ .

**196 200 Remark 2.** Although the appearance of  $\mathbb{E}[Q_m]$  may seem to be unexpected at first glance, it is usually desirable to control the concentration of a random variable around its expectation. Fortunately, it has been demonstrated in [\[38\]](#page-10-3) that for  $m = o(N^{2/5})$ , the difference  $\mathbb{E}[Q_m] - \mathbb{E}[Q'_m]$  is bounded It has been demonstrated in [38] that for  $m = o(N^{-1})$ , the difference  $\mathbb{E}[Q_m] - \mathbb{E}[Q_m]$  is bounded<br>by  $\sqrt{m}$ . Consequently, our theorems can be employed to effectively manage the deviations of  $Q'_m$ from its expectation  $\mathbb{E}[Q'_m]$  at a fast rate.

**202 203 204 205 206 207 208 209 210** In fact, we draw inspiration from the proof presented in [\[38\]](#page-10-3) and use Hoeffding's reduction method to build the connection between the sequences of random variables sampling with and without replacement. However, extending the results to the classes of sub-Gaussian and sub-exponential functions presents challenges. On one hand, the classical truncation technique yields tail bounds, nonetheless we need to combine the sequences of random variables sampling with and without replacement using moment generating functions while ensuring their convexity. This is crucial as Hoeffding's reduction method requires convexity. On the other hand, the introduction of the unbounded assumption introduces an additional term, which complicates the construction of convex moment generating functions (MGF) and the application of cheronff's method.

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## <span id="page-3-1"></span>4 GENERALIZATION BOUNDS FOR TRANSDUCTIVE LEARNING

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**214 215** Our concentration inequalities have broad applications and can serve as an important tool in learning theory when considering sampling without replacement for classes of sub-exponential functions. In this section, we will provide two examples to illustrate the risk bounds in transductive learning.

### <span id="page-4-0"></span>**216 217 218** 4.1 FAST EXCESS RISK BOUNDS FOR TRANSDUCTIVE LEARNING WITH UNBOUNDED LOSSES

**219 220 221 222** We apply our newly concentration inequalities to give fast excess risk bounds for transductive learning on ERM with unbounded losses, which is, to the best of our knowledge, the first results. We mainly follows the traditional technique called "local Rademacher complexity" developed by [\[2\]](#page-9-8). We introduced the definition of Rademacher complexity for completeness.

**223 224** Definition 2 (Rademacher complexity [\[44\]](#page-11-4)). *For a function class* F *that consists of mappings from* Z *to* R*, define*

$$
\mathfrak{R}\mathcal{F} := \mathbb{E}_{\mathbf{x},v} \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n v_i f(\mathbf{x}_i) \quad and \quad \mathfrak{R}_n\mathcal{F} := \mathbb{E}_v \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n v_i f(\mathbf{x}_i),
$$

**227 228 229** *as the Rademacher complexity and the empirical Rademacher complexity of* F*, respectively, where*  $\{v_i\}_{i=1}^n$  are *i.i.d.* Rademacher variables for which  $\mathbb{P}(v_i = 1) = \mathbb{P}(v_i = -1) = \frac{1}{2}$ .

**230 231** Since Rademacher complexity could be bounded by a computable covering number of  $\mathcal F$  via Dudley's integral bound [\[35\]](#page-10-5), we give the definition of covering number for completeness as well.

**232 233 234 Definition 3** (Covering number [\[44\]](#page-11-4)). *Assume*  $(M, \text{metr}(\cdot, \cdot))$  *is a metric space, and*  $\mathcal{F} \subseteq \mathcal{M}$ *. The* ε*-convering number of the set* F *with respect to a metric* metr(·, ·) *is the size of its smallest* ε*-net cover:*

$$
\mathcal{N}(\varepsilon,\mathcal{F},\text{metr}) = \min\{m:\exists f_1,\ldots,f_m\in\mathcal{F} \text{ such that } \mathcal{F}\subseteq \cup_{j=1}^m \mathcal{B}(f_j,\varepsilon)\},
$$

$$
237 \quad \text{where } \mathcal{B}(f,\varepsilon) := \{\tilde{f} : \text{metr}(\tilde{f},f) \leq \varepsilon\}.
$$

**238** To calculate the covering number, we also need the following assumption.

<span id="page-4-1"></span>**239 240 241 Assumption 1** (Entropy bounds). *The parameter class W is separable and there exist*  $C \geq 1, K \geq 1$ *such that*  $\forall \varepsilon \in (0, K]$ *, the*  $L_2(\mathbb{P})$ *-covering numbers and the universal metric entropies of*  $\mathcal G$  *are bounded as*  $\log \mathcal{N}(\varepsilon, \mathcal{G}, L_2(\mathbb{P})) \leq \mathcal{C} \log(K/\varepsilon)$ .

**242 243 244 245** Remark 3. Assumption [1](#page-4-1) was widely adopted in fast learning rates in statistic learning [\[31,](#page-10-6) [30,](#page-10-7) [11\]](#page-9-9). In fact, if W has finite VC-dimension, then Assumption [1](#page-4-1) is satisfied [\[3,](#page-9-7) [6\]](#page-9-10). Some literature such as [\[23\]](#page-10-8) assume that the envelope function is sub-exponential, which is a much stronger assumption.

**246 247** It will be convenient to introduce the following operators, mapping functions f defined on  $X_N$  to  $\mathbb{R}$ :

$$
f_{\rm{max}}
$$

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$$
Ef = \frac{1}{N} \sum_{i=1}^{N} f(\mathbf{x}_i), \mathbf{x}_i \in \mathbf{X}_N.
$$

**250 251 252** Assume that there is a function  $\mathbf{w}_N^* \in \mathcal{W}$  satisfying  $R_N(\mathbf{w}_N^*) = \inf_{\mathbf{w} \in \mathcal{W}} R_N(\mathbf{w})$ . Define the excess loss class  $\mathcal{F}^* = \{f : f(\mathbf{x}) = \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}_N^*, \mathbf{x}), \mathbf{w} \in \mathcal{W}\}.$ 

<span id="page-4-2"></span>**Theorem 3.** Assume that there is a constant  $B > 0$  such that for every  $f \in \mathcal{F}^*$  we have  $Ef^2 \leq$ B · Ef*. Suppose Assumptions [1](#page-4-1) hold and the objective function* ℓ(·; ·) *is sub-Gaussian. For any*  $\delta \in (0, 1)$ *, with probability*  $1 - \delta$ *,* 

$$
\varepsilon_u(\hat{\mathbf{w}}_m) = \mathcal{O}\Bigg(\frac{N}{mu}\Big(\log m + \log u + \frac{N\log\frac{1}{\delta}}{m} + \frac{N\log\frac{1}{\delta}}{u} + \sqrt{\log N\log\frac{1}{\delta}}\Big)\Bigg).
$$

<span id="page-4-3"></span>**Theorem 4.** Assume that there is a constant  $B > 0$  such that for every  $f \in \mathcal{F}^*$  we have  $Ef^2 \leq$ B · Ef*. Suppose Assumptions [1](#page-4-1) hold and the objective function* ℓ(·; ·) *is sub-exponential. For any*  $\delta \in (0, 1)$ *, with probability*  $1 - \delta$ *,* 

$$
\varepsilon_u(\hat{\mathbf{w}}_m) = \mathcal{O}\Bigg(\frac{N}{mu}\Big(\log m + \log u + \frac{N\log\frac{1}{\delta}}{m} + \frac{N\log\frac{1}{\delta}}{u} + \sqrt{\log^2 N\log\frac{1}{\delta}}\Big)\Bigg).
$$

**264 265 266 267 268 269** Remark 4. By utilizing variance information and introducing the Bernstein condition, we present the first results for fast learning rates under unbounded losses. Applying our concentration inequalities under unbounded conditions to local Rademacher method is not a straightforward task. We need to skillfully separate variance term and the Orlicz norm term through inequalities while constructing a suitable partition. Similarly, when employing the localized approach, we need to create a slightly modified version for partition  $E_m f$  which is affected by the Hoeffding's reduction method applied during the proof of our concentration inequalities given in Section [3.](#page-2-0)

### <span id="page-5-0"></span>**270 271** 4.2 IMPROVED BOUNDS OF GNNS WITH SGD

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**272 273 274 275** GNNs have achieved great success in practice, but research on the generalization performance of GNNs for node classification remains limited. In the real world, training nodes are sampled without replacement from the entire node set, and test nodes remain visible during training [\[13,](#page-9-4) [32\]](#page-10-9), which perfectly fits the transductive learning setting.

**276 277 278 279** The current state-of-the-art work on generalization error for graph node classification [\[37\]](#page-10-4) was based on the concentration inequality for transductive learning provided by [\[13\]](#page-9-4). In this subsection, we aim to obtain a tighter generalization upper bound by applying our new concentration inequalities introduced in this paper.

**280 281 282 283** Let's introduce some notations for GNNs firstly. Consider an undirected graph  $\mathcal{G} = \{ \mathcal{V}, \mathcal{E} \}$ , where  $V$  represents a set of nodes and  $E$  represents the edges between these nodes. The graph has a total of  $n = |\mathcal{V}|$  nodes. Each node corresponds to an instance denoted as  $z_i = (x_i, y_i)$ , comprising a feature vector  $x_i$  and a label  $y_i$  from a space  $\mathcal{Z} = \mathcal{X} \times \mathcal{Y}$ .

**284 285 286 287 288 289 290 291 292 293 294 295** Let **X** denote the feature matrix, where the *i*-th row  $X_{i*}$  represents the feature  $x_i$ . The adjacency matrix is represented as A, and the diagonal degree matrix is denoted as D. Specifically, the diagonal entry  $D_{ii}$  is computed as the sum of the weights of the edges connected to node i. We introduce the normalized adjacency matrix  $\tilde{A} = (D+I_n)^{-\frac{1}{2}}(A+I_n)(D+I_n)^{-\frac{1}{2}}$ , where  $I_n$  is the identity matrix of size  $n \times n$ , and  $\sqrt{|\mathcal{Y}|}$  corresponds to the square root of the number of categories. This matrix accounts for self-loops and captures the graph's normalized connectivity structure, aiding in subsequent analyses. We limit the scope of the learner to a given GNN and let w be its learnable parameters. Given the isomorphism between  $\mathbb{R}^{p \times q}$  and  $\mathbb{R}^{pq}$ , our analysis in this work focuses on the more concise vector space. To achieve this, we introduce a unified vector  $\mathbf{w} = [\text{vec}[\mathbf{W}_1]; \dots; \text{vec}[\mathbf{W}_H]]$ to represent the collection  ${ \bf \{W}_h \}_{h=1}^H,$  where  ${\rm vec}[\cdot]$  denotes the vectorization operator that transforms a given matrix into a vector. In other words,  $vec[\mathbf{W}] = [\mathbf{W}_{*1}; \dots; \mathbf{W}_{*q}]$  for  $\mathbf{W} \in \mathbb{R}^{p \times q}$ . In this context,  $W_{*i}$  represents the *i*-th column of W.

**296 297 298 299 300 301** In this section, we apply the concentration inequalities presented in this paper to derive improved rates of the current optimal results [\[37\]](#page-10-4) for GNNs with SGD (Algorithm [1\)](#page-6-0). The initialization weight of the model is denoted as  $w^{(1)}$ . We use  $b_g$  to represent the supremum of the gradient when evaluated at the initialized parameters, defined as  $b_g = \sup_{z \in \mathcal{Z}} ||\nabla \ell(\mathbf{w}^{(1)}; z)||_2$ . The activation function is represented by  $\omega(\cdot)$ .

**302 303 304** We notice that since the full data  $X_N$  is given, then  $R_N(\mathbf{w}) = \frac{1}{N} \sum_{i=1}^N \ell(\mathbf{w}; \mathbf{x}_i)$  is not a random variable. Also, for any training sample  $X_m$ , the test error  $R_u(\mathbf{w})$  can be expressed in terms of  $R_N(\mathbf{w})$  and the training error  $R_m(\mathbf{w})$  as follows:

$$
R_u(\mathbf{w}) = \frac{1}{u} \sum_{i=m+1}^{m+u} \ell(\mathbf{w}; \mathbf{x}_i) = \frac{1}{u} \left( (m+u) R_N(\mathbf{w}) - \sum_{i=1}^m \ell(\mathbf{w}; \mathbf{x}_i) \right) = \frac{m+u}{u} R_N(\mathbf{w}) - \frac{m}{u} \hat{R}_m(\mathbf{w}).
$$

**308 310** Thus, for any fixed  $\mathbf{w} \in \mathcal{W}$ , the quantity  $R_u(\mathbf{w}) - \hat{R}_m(\mathbf{w}) = \frac{N}{u}(R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w}))$ , for any  $\hat{\mathbf{w}}$ , we have

$$
R_u(\hat{\mathbf{w}}) - \hat{R}_m(\hat{\mathbf{w}}) \le \sup_{\mathbf{w} \in \mathcal{W}} R_u(\mathbf{w}) - \hat{R}_m(\mathbf{w}) = \frac{N}{u} \sup_{\mathbf{w} \in \mathcal{W}} R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w}).
$$

**312 313 314 315 316** Note that for any fixed  $\mathbf{w} \in \mathcal{W}$ ,  $\mathbb{E}_{\mathbf{x}}[R_N(\mathbf{w}) - \ell(\mathbf{w}; \mathbf{x})] = R_N(\mathbf{w}) - \mathbb{E}_{\mathbf{x}}\ell(\mathbf{w}; \mathbf{x}) = 0$ , thus, we can use the transductive setting described in Section [3.](#page-2-0) Considering the function class  $\mathcal{F}_{w} := \{f_{w} :$  $f_{\mathbf{w}}(\mathbf{x}) = R_N(\mathbf{w}) - \ell(\mathbf{w}; \mathbf{x}), \mathbf{w} \in \mathcal{W}$  associated with W. For fixed w,  $R_N(\mathbf{w})$  is not random, at the same time, centering random variable does not change its variance, so we have

$$
\sigma_{\mathcal{W}}^2 = \sup_{f_{\mathbf{w}} \in \mathcal{F}_{\mathbf{w}}} \mathbb{V}[f_{\mathbf{w}}(\mathbf{x})] = \sup_{\mathbf{w} \in \mathcal{W}} \mathbb{V}[\ell(\mathbf{w}; \mathbf{x})] = \sup_{\mathbf{w} \in \mathcal{W}} \left( \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} (\ell(\mathbf{w}; \mathbf{x}) - R_N(\mathbf{w}))^2 \right).
$$

**320 321** Using Theorem [1](#page-2-2) and [2,](#page-3-0) we can obtain the results that hold without any other assumptions, expect for the classes of sub-Gaussian or sub-exponential functions on the learning problem

322  
\n
$$
\sup_{\mathbf{w}\in\mathcal{W}}(R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w})) \le (1+\eta)E_m + 2\sqrt{\left(\frac{4(1+\beta)\sigma_W^2}{m} + \frac{2C^2\|\max_{\mathbf{x}}\sup_{f\in\mathcal{F}_{\mathbf{w}}f_{\mathbf{w}}(\mathbf{x})\|_{\psi_2}^2}{m^2}\right)\log\frac{6}{\delta}},
$$

**324**



<span id="page-6-0"></span>

and

$$
\sup_{\mathbf{w}\in\mathcal{W}}(R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w})) \leq (1+\eta)E_m + 4\sqrt{\left(\frac{(1+\beta)\sigma_W^2}{m} + \frac{3C^2\|\max_\mathbf{x}\sup_{f\in\mathcal{F}_\mathbf{w}}f_\mathbf{w}(\mathbf{x})\|_{\psi_1}^2}{m^2}\right)\log\frac{6}{\delta}},
$$

where let  $\{\xi_1,\ldots,\xi_n\}$  be random variables sampled with replacement from  $X_N$  and denote

$$
E_m = \mathbb{E}\left[\sup_{\mathbf{w}\in\mathcal{W}}\left(R_N(\mathbf{w}) - \frac{1}{m}\sum_{i=1}^m \ell(\mathbf{w};\xi_i)\right)\right].
$$

<span id="page-6-1"></span>Next, we need to derive the upper bounds of  $\sigma_W^2$ ,  $E_m$  and  $\|\max_{\mathbf{x}} \sup_{f \in \mathcal{F}_{\mathbf{w}}} f_{\mathbf{w}}(\mathbf{x})\|_{\psi_\alpha}^2$ ,  $\alpha = 1$  or 2 in GNNs with SGD. We present the assumptions only used in this subsection.

**346 347 Assumption 2.** Assume that there exists a constant  $c_X > 0$  such that  $\|\mathbf{x}\|_2 \le c_X$  holds for all  $\mathbf{x} \in \mathcal{X}$  *and there exists a constant*  $c_W > 0$  *such that*  $||\mathbf{W}_h|| \leq c_W$ ,  $h \in [H]$  *for*  $\mathbf{w} \in \mathcal{W}$ *.* 

**348 349 350 351 352** Remark 5. Assumption [2](#page-6-1) necessitates boundness of input features as discussed by [\[42\]](#page-11-5) and the boundness of parameters during the training process, which is a common consideration in the generalization analysis of Graph Neural Networks (GNNs) [\[16,](#page-9-11) [28,](#page-10-10) [9,](#page-9-12) [15\]](#page-9-13). This assumption play a crucial role in the analysis of Lipschitz continuity and Hölder smoothness of the objective with respect to the parameters w.

<span id="page-6-2"></span>**353 354 Assumption 3.** Assume that the activation function  $\omega(\cdot)$  is  $\tilde{\alpha}$ -Höder smooth. To be specific, let  $P > 0$  and  $\tilde{\alpha} \in (0, 1]$ , for all  $\mathbf{u}, \mathbf{v} \in \mathbb{R}^d$ ,

$$
\frac{355}{356}
$$

 $\|\nabla\omega(\mathbf{u})-\nabla\omega(\mathbf{v})\|_2 \leq P\|\mathbf{u}-\mathbf{v}\|_2^{\tilde{\alpha}}.$ 

**357 358 359 360 361 362** Remark 6. It can be established that Assumption [3](#page-6-2) leads to the Lipschitz continuity of the activation function when  $\tilde{\alpha} = 0$ . Furthermore,  $\tilde{\alpha} = 1$  implies the smoothness of the activation function. As a result, Assumption [3](#page-6-2) stands as notably milder in comparison to the assumption found in prior works [\[42,](#page-11-5) [9\]](#page-9-12), which mandates the activation function's smoothness. In order to facilitate analysis without introducing a significant disparity between theory and practical application, we often use modified ReLU function

$$
\frac{363}{364}
$$

**365 366**

**376 377**

$$
\omega(x) = \begin{cases} 0, x \le 0, \\ x^q, 0 < x \le \left(\frac{1}{q}\right)^{\frac{1}{q-1}}, \\ x - \left(\frac{1}{q}\right)^{\frac{1}{q-1}} + \left(\frac{1}{q}\right)^{\frac{q}{q-1}}, x > \left(\frac{1}{q}\right)^{\frac{1}{q-1}}. \end{cases}
$$

This modified function, controlled by the hyperparameter  $q \in (1, 2]$ , not only satisfies Assumption [3](#page-6-2) but also maintains an acceptable approximation to the vanilla ReLU function.

<span id="page-6-3"></span>Lemma 1 (Proposition 4.1 in [\[37\]](#page-10-4)). *Suppose that Assumption [2](#page-6-1) and [3](#page-6-2) hold. Denote by* F *a specific GNN, for any*  $w, w' \in W$  *and*  $x \in X_N$ *, the objective*  $\ell(w; x)$  *satisfies* 

$$
|\ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}'; \mathbf{x})| \leq L_{\mathcal{F}} \|\mathbf{w} - \mathbf{w}'\|_2,
$$

**374 375** *and*

$$
\|\nabla \ell(\mathbf{w}; \mathbf{x}) - \nabla \ell(\mathbf{w}'; \mathbf{x})\| \leq P_{\mathcal{F}} \max\{\|\mathbf{w} - \mathbf{w}'\|_2^{\tilde{\alpha}}, \|\mathbf{w} - \mathbf{w}'\|_2\},\
$$

*with constant*  $L_F$  *and*  $P_F$ *.* 

**378 379 380 381 382** Remark 7. [\[37\]](#page-10-4) demonstrates that several widely used structured networks in GNNs such as GCN [\[20\]](#page-10-11), GCNII [\[7\]](#page-9-14), SGC [\[45\]](#page-11-6), APPNP [\[17\]](#page-9-15) and GPR-GNN [\[8\]](#page-9-16) satisfy Lemma [1.](#page-6-3) We leverage the properties of these network structures in Lemma [1](#page-6-3) to derive improved upper bounds using our concentration inequalities instead of [\[13\]](#page-9-4).

**383** The following two assumptions are introduced to obtain the optimization error.

<span id="page-7-0"></span>**384 Assumption 4.** Assume that there exist a constant  $G > 0$  such that for all  $\mathbf{x} \in \mathbf{Z}$ 

 $\sqrt{\eta_t} \|\nabla \ell(\mathbf{w}_t; \mathbf{x})\|_2 \leq G$ 

*holds for any*  $t \in \mathbb{N}$ , where  $\{\eta_t\}_{t=1}^T$  is learning rates.

<span id="page-7-1"></span>**Assumption 5.** Assume that there exists a constant  $\sigma_0 > 0$  such that for  $\forall t \in \mathbb{N}_+$ , the following *inequality holds*

 $\mathbb{E}_{jt}[\|\nabla \ell(\mathbf{w}; \mathbf{x}_{jt})\|^2] \leq \sigma_0^2.$ 

**392 393 394 395 396 Remark 8.** Assumption [4](#page-7-0) [\[26,](#page-10-12) [27\]](#page-10-13) requires a bound on the product of the gradient and the square root of the step sizes. This condition is weaker than the commonly employed bounded gradient assumption [\[18,](#page-9-17) [21\]](#page-10-14), as the learning rate naturally approaches zero throughout the iteration process. Assumption [5](#page-7-1) requires the boundness of variances of stochastic gradients, which is a standard assumption in stochastic optimization studies [\[21,](#page-10-14) [26,](#page-10-12) [27\]](#page-10-13).

**397** Now, we can derive the risk bounds of GNNs with SGD.

<span id="page-7-2"></span>**399 Theorem [5](#page-7-1).** *Suppose Assumptions* [2,](#page-6-1) [3,](#page-6-2) [4,](#page-7-0) and 5 hold, and assume the objective function  $\ell(\cdot;\cdot)$  be  $s$ ub-Gaussian. Suppose that the step sizes  $\{\eta_t\}$  satisfies  $\eta_t = \frac{1}{t+t_0}$  such that  $t_0 \ge \max\{(2P)^{1/\alpha}, 1\}.$ *For any*  $\delta \in (0, 1)$ *, with probability*  $1 - \delta$ *,* 

(a). If 
$$
\alpha \in (0, \frac{1}{2})
$$
, we have

$$
R_u(\mathbf{w}_1^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\left(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)T^{\frac{1-2\alpha}{2}}\log\left(\frac{1}{\delta}\right) + \frac{N\log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right)
$$

.

(*b*). If  $\alpha = \frac{1}{2}$ , we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log(T)\log\left(\frac{1}{\delta}\right) + \frac{N\log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\bigg).
$$

*(c). If*  $\alpha \in (\frac{1}{2}, 1]$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)\log\bigg(\frac{1}{\delta}\bigg) + \frac{N\log\big(\frac{1}{\delta}\big)}{u\sqrt{m}}\bigg).
$$

**416 417 418 419 420 421 422 423 424** Remark 9. Similar result for sub-exponential loss functions is given in Appendix C. Generally, **Remark 9.** Similar result for sub-exponential loss functions is given in Appendix C. Generally, comparing our bound with [\[37\]](#page-10-4), their bound is of order  $\mathcal{O}\left(\left(\frac{1}{m}+\frac{1}{u}\right)\sqrt{m+u}\right)$  after the  $L_{\mathcal{F}}$  but our bounds are of order  $\mathcal{O}\left(\frac{\sqrt{m+u}}{u}\right)$ , at the same time, we have an extra term  $\frac{m+u}{u\sqrt{m}}$ , which is introduced due to the variance information. Notice that it's not as if they didn't have the second term, because their first term is larger than the second one and so the final magnitude doesn't change. Our results are better when  $m = o(N^{2/5})$ . We can take a more visual example to demonstrate the advantages of our bounds. For  $m = \Theta(N^{1/5})$ , our bound is of order  $\mathcal{O}\left(\frac{1}{\sqrt{m}}\right)$  but their bound is of order  $\mathcal{O}(m^3)$ , which fails to provide a reasonable generalization guarantee.

**426 427** Similarly, we can also derive a upper bound of the test error under PL condition following proof trajectory of [\[37\]](#page-10-4).

**Assumption 6** (PL-condition). *Suppose that there exists a constant*  $\mu$  *such that for all*  $\mathbf{w} \in \mathcal{W}$ ,

$$
\hat{R}_{m}(\mathbf{w}) - \hat{R}_{m}(\hat{\mathbf{w}}^*) \leq \frac{1}{2\mu} \|\nabla \hat{R}_{m}(\mathbf{w})\|_2,
$$

*holds for the given set*  $X_m$  *from*  $X_N$ *.* 

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<span id="page-8-0"></span>**432 433 434 435 Corollary 1.** Suppose Assumptions [2,](#page-6-1) [3,](#page-6-2) [4,](#page-7-0) and [5](#page-7-1) hold and assume the objective function  $\ell(\cdot;\cdot)$ *be sub-Gaussian.* Suppose that the learning rate  $\{\eta_t\}$  satisfies  $\eta_t = \frac{2}{\mu(t+t_0)}$  such that  $t_0 \ge$  $\max\{\frac{2}{\mu}(2P)^{\frac{1}{\alpha}},1\}$ *. For any*  $\delta \in (0,1)$ *, with probability*  $1-\delta$ *,* 

*(a). If*  $\alpha \in (0, \frac{1}{2})$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)T^{\frac{1}{2}-\alpha}\log\bigg(\frac{1}{\delta}\bigg) + \frac{N\log\big(\frac{1}{\delta}\big)}{u\sqrt{m}} + \frac{1}{T^{\alpha}}\bigg),
$$

(*b*). If  $\alpha = \frac{1}{2}$ , we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log(T)\log\left(\frac{1}{\delta}\right) + \frac{N\log\left(\frac{1}{\delta}\right)}{u\sqrt{m}} + \frac{1}{T^{\alpha}}\bigg).
$$

*(c). If*  $\alpha \in (\frac{1}{2}, 1)$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)\log(1/\delta) + \frac{N\log\left(\frac{1}{\delta}\right)}{u\sqrt{m}} + \frac{1}{T^{\alpha}}\bigg).
$$

*(d). If*  $\alpha = 1$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - R_u(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)\log(1/\delta) + \frac{N\log\left(\frac{1}{\delta}\right)}{u\sqrt{m}} + \frac{\log(T)\log^3(1/\delta)}{T}\bigg).
$$

**458** Remark 10. For completeness, we present Corollary [1](#page-8-0) for sub-Gaussian and Corollary [2](#page-30-0) (See Appendix [C.2\)](#page-29-0) for sub-exponential. There is nothing special about the proofs, which simply combine Theorem [5](#page-7-2) and Theorem [11](#page-29-1) with existing optimization results. The results under the sub-exponential distribution are provided in Appendix [4.2.](#page-5-0) It is worth point out that all the popular neural network structures introduced in [\[37\]](#page-10-4) can be applied to our results to obtain bounds that make sense.

**463 464 465 466 467** Our work in this section differs significantly from that of [\[37\]](#page-10-4). They used the concentration inequalities based on [\[13\]](#page-9-4) to derive generalization bounds, while proving that certain modern neural network structures satisfy Lipschitz continuity under their assumptions. In contrast, we employ newly proposed concentration inequalities that relax the boundness condition and also consider variance information which obtain improved rates under the same settings.

**468 469 470 471 472 473 474 475 476 477** While previous papers have utilized technologies based on concentration inequalities proposed by [\[13\]](#page-9-4) and then bound the transductive Rademacher complexity, deriving the generalization error using our new inequality is not straightforward. We need to derive the upper bounds for  $\sigma_w^2$ ,  $E_m$ , and  $\|\max_{\mathbf{x}} \sup_{f \in \mathcal{F}_{\mathbf{w}}} f_{\mathbf{w}}(\mathbf{x})\|_{\psi_\alpha}^2$ , respectively.  $\sigma_{\mathbf{w}}^2$  needs to be bounded using concentration inequalities for unbounded distributions. For the sub-exponential distribution, we even need to introduce the concentration inequalities under the sub-Weibull distribution to address the issue.  $E_m$  is introduced due to the Hoeffding's reduction method and is distinct from the traditional gap between the population and the samples. This requires us to convert it into Rademacher complexity and then use the covering number to obtain the upper bound. The term  $\|\max_{\mathbf{x}} \sup_{f \in \mathcal{F}_{\mathbf{w}}} f_{\mathbf{w}}(\mathbf{x})\|_{\psi_{\alpha}}^2$  is introduced due to the unbounded assumption. We utilize pisier's inequality [\[34\]](#page-10-15) to present the max operator before the Orlicz norm.

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## 5 CONCLUSION

**482 483 484 485** In this paper, we focus on transductive learning settings. Firstly, we introduce two newly concentration inequalities for the suprema of empirical processes sampled without replacement for unbounded functions. Using our inequalities, we derive the first fast risk bounds for ERM in transductive learning under bounded losses. On the other hand, we provide improved risk bounds for GNNs with SGD, which is better than the state-of-the-art work [\[37\]](#page-10-4) when  $m = o(N^{2/5})$ .

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<span id="page-11-6"></span><span id="page-11-5"></span><span id="page-11-4"></span><span id="page-11-3"></span><span id="page-11-2"></span><span id="page-11-1"></span> 

# A ADDITIONAL DEFINITIONS AND LEMMATA

<span id="page-12-2"></span>**Theorem 6** ([\[19\]](#page-10-16)). Let  $\{U_1, \ldots, U_m\}$  and  $\{W_1, \ldots, W_m\}$  be sampled uniformly from a finite set of d-dimensional vectors  $\{v_1,\ldots,v_N\}\subset\mathbb{R}^d$  with and without replacement, respectively. Then, for any continuous and convex function  $F: \mathbb{R}^d \to \mathbb{R}$ , the following holds:

$$
\mathbb{E}\left[F\left(\sum_{i=1}^m W_i\right)\right] \leq \mathbb{E}\left[F\left(\sum_{i=1}^m U_i\right)\right].
$$

<span id="page-12-1"></span>**Lemma 2** ([\[38\]](#page-10-3)). Let  $\boldsymbol{x} = (x_1, \ldots, x_d)^T \in \mathbb{R}^d$ . Then the following function is convex for all  $\lambda > 0$ 

$$
F(\boldsymbol{x}) = \exp\left(\lambda \sup_{i=1,\dots,d} x_i\right).
$$

<span id="page-12-3"></span>Theorem 7 (Theorem 4 via Pisier's inequality [\[34\]](#page-10-15)). *For independent real random variables*  $Y_i, \ldots, Y_n$ , we have the following inequality that

$$
\left\|\max_{i\leq n}Y_i\right\|_{\psi_{\alpha}}\leq K_{\alpha}\max_{i\leq n}\|Y_i\|_{\psi_{\alpha}}\log^{1/\alpha} n,
$$

*where*  $K_{\alpha}$  *is a positive constant.* 

Definition 4 (Rademacher complexity [\[44\]](#page-11-4)). *For a function class* F *that consists of mappings from* Z *to* R*, define*

$$
\mathfrak{R}\mathcal{F} := \mathbb{E}_{\mathbf{x},v} \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n v_i f(\mathbf{x}_i) \quad and \quad \mathfrak{R}_n\mathcal{F} := \mathbb{E}_v \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n v_i f(\mathbf{x}_i),
$$

**672 673 674** *as the Rademacher complexity and the empirical Rademacher complexity of* F*, respectively, where*  $\{v_i\}_{i=1}^n$  are i.i.d. Rademacher variables for which  $\mathbb{P}(v_i = 1) = \mathbb{P}(v_i = -1) = \frac{1}{2}$ .

**Definition 5** (Covering number [\[44\]](#page-11-4)). *Assume*  $(M, \text{metr}(\cdot, \cdot))$  *is a metric space, and*  $\mathcal{F} \subseteq \mathcal{M}$ *. The* ε*-convering number of the set* F *with respect to a metric* metr(·, ·) *is the size of its smallest* ε*-net cover:*

$$
\mathcal{N}(\varepsilon,\mathcal{F},\text{metr}) = \min\{m:\exists f_1,\ldots,f_m\in\mathcal{F} \text{ such that } \mathcal{F}\subseteq \cup_{j=1}^m \mathcal{B}(f_j,\varepsilon)\},
$$

*where*  $\mathcal{B}(f, \varepsilon) := \{ \tilde{f} : \text{metr}(\tilde{f}, f) \leq \varepsilon \}.$ 

Lemma 3 (Dudley's integral bound [\[35\]](#page-10-5)). *Given* r > 0 *and class* F *that consists of functions defined on* Z*,*

**683**

$$
\Re_n\{f \in \mathcal{F} : \mathbb{P}_n[f^2] \le r\} \le \inf_{\varepsilon_0 > 0} \left\{ 4\varepsilon_0 + 12 \int_{\varepsilon_0}^{\sqrt{r}} \sqrt{\frac{\log \mathcal{N}(\varepsilon, \mathcal{F}, L_2(\mathbb{P}_n))}{n}} d\varepsilon \right\}
$$

.

Definition 6 ([\[43\]](#page-11-3)). *A random variable* X *is sub-Weibull random variables with taill parameter* θ *when for any*  $x > 0$ ,

$$
\mathbb{P}(X \ge x) = \exp(-bx^{1/\theta}), \text{ for some } b > 0, \theta > 0.
$$

<span id="page-12-4"></span>**Lemma 4.** (Concentration of the sum for sub-Weibull distribution [\[43\]](#page-11-3)) Let that  $X_1, \ldots, X_n$  be identically distributed sub-Weibull random variables with tail parameter  $\theta$ . Then, for all  $x \geq nK_{\theta}$ , we have

$$
\mathbb{P}\left(\left|\sum_{i=1}^n X_i\right| \geq x\right) \leq \exp\left(-\left(\frac{x}{nK_{\theta}}\right)^{1/\theta}\right),\,
$$

**695 696** for some constant  $K_{\theta}$  dependent on  $\theta$ .

<span id="page-12-0"></span>**Theorem 8** ([\[1\]](#page-9-2)). Let  $X_1, \ldots, X_m$  be independent random variables with values in a measurable *space* (S, B) *and let* F *be a countable class of measurable functions*  $f : S \to [-a, a]$ *, such that for all* i,  $\mathbb{E} f(X_i) = 0$ . Consider the random variable

**699**

$$
Q = \sup_{f \in \mathcal{F}} \sum_{i=1}^{m} f(X_i)
$$

**702 703** *and*

**704 705**

**717 718 719**

$$
\sigma^2 = \sup_{f \in \mathcal{F}} \mathbb{E} f(X_1)^2.
$$

*Then, for all*  $0 < \eta \leq 1$ ,  $\beta > 0$  *there exists a constant*  $C = C(\eta, \beta)$ *, such that for all t* > 0*,* 

$$
\mathbb{P}(Q - (1 + \eta)\mathbb{E}Q \ge t) \le \exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + \exp\left(-\frac{t}{Ca}\right)
$$

,

.

*and*

$$
\mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t) \le \exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + \exp\left(-\frac{t}{Ca}\right).
$$

<span id="page-13-0"></span>**713 714 715 716** Theorem 9. (Tail inequality for suprema of empirical process corresponding to classes of sub-**Gaussian functions**) Let  $X_1, \ldots, X_m$  be independent random variables with values in a measurable space (S, B) and let F be a countable class of measurable functions  $f : S \to \mathbb{R}$ . Assume that for every  $\hat{f} \in \mathcal{F}$  and every  $i$ ,  $\mathbb{E}f(X_i) = 0$  and  $|| \sup_f |f(X_i)||_{\psi_2} < \infty$ . Let

$$
Q = \sup_{f \in \mathcal{F}} \sum_{i=1}^{m} f(X_i)
$$

**720** and

$$
\sigma^2 = \sup_{f \in \mathcal{F}} \mathbb{E} f(X_i)^2.
$$

Then, for all  $0 < \eta < 1$  and  $\beta > 0$ , there exists a constant  $C = C(\eta, \beta)$ , such that for all  $epsilon > 0$ ,

$$
\mathbb{P}(Q - (1+\eta)\mathbb{E}Q \ge t) \le \exp\left(-\frac{t^2}{2(1+\beta)m\sigma^2}\right) + 3\exp\left(-\left(\frac{t}{C\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_2}}\right)^2\right),
$$

and

$$
\mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t) \le \exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + 3\exp\left(-\left(\frac{t}{C\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_2}}\right)^2\right)
$$

<span id="page-13-1"></span>Theorem 10. (Tail inequality for suprema of empirical process corresponding to classes of sub-exponential functions) Let  $X_1, \ldots, X_m$  be independent random variables with values in a measurable space (S, B) and let F be a countable class of measurable functions  $f : \mathcal{S} \to \mathbb{R}$ . Assume that for every  $f \in \mathcal{F}$  and every  $i$ ,  $\mathbb{E}f(X_i) = 0$  and  $|| \sup_f |f(X_i)||_{\psi_1} < \infty$ . Let

$$
Q = \sup_{f \in \mathcal{F}} \sum_{i=1}^{m} f(X_i)
$$

and

$$
\sigma^2 = \sup_{f \in \mathcal{F}} \mathbb{E} f(X_i)^2.
$$

**743 744** Then, for all  $0 < \eta < 1$  and  $\beta > 0$ , there exists a constant  $C = C(\eta, \beta)$ , such that for all  $epsilon > 0$ ,

$$
\mathbb{P}(Q - (1+\eta)\mathbb{E}Q \ge t) \le \exp\left(-\frac{t^2}{2(1+\beta)m\sigma^2}\right) + 3\exp\left(-\frac{t}{C\|\max_i \sup_{f \in \mathcal{F}}f(X_i)\|_{\psi_1}}\right),
$$

**748** and

$$
\mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t) \le \exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + 3\exp\left(-\frac{t}{C\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_1}}\right).
$$

**752 753 754 755** The proofs of Theorem [9](#page-13-0) and Theorem [10](#page-13-1) are similar with [\[1\]](#page-9-2), which under the assumption that the summands have finite  $\psi_{\alpha}$  Orlicz norm with  $\alpha \in (0,1)$  and they analyze the random variable  $Q = \sup_{f \in \mathcal{F}} |\sum_{i=1}^m f(X_i)|$ . However, in this paper, we consider  $Q = \sup_{f \in \mathcal{F}} \sum_{i=1}^m f(X_i)$ . In consequence we give the sub-gaussian and sub-exponential version ( $\alpha = 1, 2$ ) for the sake of completeness here.

**745 746 747**

**756 757** *Proof of Theorem [9](#page-13-0) and Theorem [10.](#page-13-1)* Without loss of generality, we assume that

<span id="page-14-4"></span>
$$
t/\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f(X_i)\|_{\psi_\alpha} > K(\alpha,\eta,\beta),\tag{1}
$$

**760 761 762** otherwise we can make the theorem trivial by choosing the constant  $C = C(\alpha, \eta, \beta)$  to be large enough. The conditions on the constant  $K(\alpha, \eta, \beta)$  will be imposed later in the following proof.

**763 764 765 766 767** Let  $\varepsilon = \varepsilon(\beta) > 0$  which will be determined later and for all  $f \in \mathcal{F}$  consider the truncated functions  $f_1(x) = f(x) \mathbf{1}_{\{\sup_{f \in \mathcal{F}} |f(x)| \leq \rho\}}$  (the truncation level  $\rho$  will be determined and fixed later). Define the functions  $f_2(x) = f(x) - f_1(x) = f(x) \mathbf{1}_{\{\sup_{f \in \mathcal{F}} |f(x)| > \rho\}}$ . Let  $\mathcal{F}_i = \{f_i : f \in \mathcal{F}\}\)$ . Then we have

<span id="page-14-0"></span>
$$
Q = \sup_{f \in \mathcal{F}} \sum_{i=1}^{m} f(X_i) \le \sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^{m} (f_1(X_i) - \mathbb{E} f_1(X_i)) + \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^{m} (f_2(X_i) - \mathbb{E} f_2(X_i)), \quad (2)
$$

**771** and

**758 759**

**768 769 770**

$$
Q \geq \sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E} f_1(X_i)) - \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E} f_2(X_i)),
$$
 (3)

**776 777 778** where the above inequalities satisfy because of the fact that  $\mathbb{E} f_1(X_i) + \mathbb{E} f_2(X_i) = 0$  for all  $f \in \mathcal{F}$ . Similarly, by Jensen's inequality, we have

<span id="page-14-2"></span><span id="page-14-1"></span>
$$
\mathbb{E} \sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E} f_1(X_i)) - 2\mathbb{E} \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m f_2(X_i)
$$
\n
$$
\leq \mathbb{E} Q \leq \sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E} f_1(X_i)) + 2\mathbb{E} \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m f_2(X_i).
$$
\n(4)

Denoting

$$
A = \mathbb{E} \sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^{m} (f_1(X_i) - \mathbb{E} f_1(X_i))
$$

and

$$
B = \mathbb{E} \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m f_2(X_i).
$$

Combining [\(2\)](#page-14-0) and [\(4\)](#page-14-1), we get

<span id="page-14-3"></span>
$$
\mathbb{P}(Q - (1 + \eta)\mathbb{E}Q \ge t)
$$
\n
$$
\leq \mathbb{P}\left(\sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E}f_1(X_i)) \ge (1 + \eta)\mathbb{E}Q + (1 - \varepsilon)t\right)
$$
\n
$$
+ \mathbb{P}\left(\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E}f_2(X_i)) \ge \varepsilon t\right)
$$
\n
$$
\leq \mathbb{P}\left(\sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E}f_1(X_i)) \ge (1 + \eta)A - 4B + (1 - \varepsilon)t\right)
$$
\n
$$
+ \mathbb{P}\left(\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E}f_2(X_i)) \ge \varepsilon t\right).
$$
\n(5)

**810 811** Similarly, combing [\(3\)](#page-14-2) and [\(4\)](#page-14-1), we have

$$
\begin{array}{c} 812 \\ 813 \end{array}
$$

$$
\mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t)
$$

$$
\begin{array}{c} 814 \\ 815 \end{array}
$$

<span id="page-15-0"></span>
$$
\leq \mathbb{P}\left(\sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E}f_1(X_i)) \leq (1-\eta)\mathbb{E}Q - (1-\varepsilon)t\right) \n+ \mathbb{P}\left(\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E}f_2(X_i)) \geq \varepsilon t\right) \n\leq \mathbb{P}\left(\sup_{f_1 \in \mathcal{F}_1} \sum_{i=1}^m (f_1(X_i) - \mathbb{E}f_1(X_i)) \geq (1-\eta)A + 2B - (1-\varepsilon)t\right) \n+ \mathbb{P}\left(\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E}f_2(X_i)) \geq \varepsilon t\right).
$$
\n(6)

Next, we need to choose proper truncation level  $\rho$  in a way, which would allow to bound the first summands on the right-hand sides of [\(5\)](#page-14-3) and [\(6\)](#page-15-0) with Theorem [8.](#page-12-0)

**828** Let us set

<span id="page-15-1"></span>
$$
\rho = 8\mathbb{E} \max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i) \le K_\alpha \left\| \max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i) \right\|_{\psi_\alpha}.
$$
 (7)

Notice that by the Chebyshev inequality and the definition of the class  $\mathcal{F}_2$ , we have

$$
\mathbb{P}\left(\max_{k\leq m}\sup_{f\in\mathcal{F}}\sum_{i=0}^k f_2(X_i)>0\right)\leq \mathbb{P}\left(\max_i \sup_f f(X_i)>\rho\right)\leq 1/8.
$$

Thus by the Hoffmann-Jorgensen inequality [\[25\]](#page-10-17), we get

<span id="page-15-2"></span>
$$
B = \mathbb{E} \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m f_2(X_i) \le 8 \mathbb{E} \max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i). \tag{8}
$$

In consequence

$$
\mathbb{E}\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E}f_2(X_i)) \le 16 \mathbb{E}\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i) \le K_\alpha \left\|\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_\alpha}.
$$

Thus, we have

$$
\left\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f_2(X_i)-\mathbb{E}f_2(X_i)\right\|_{\psi_{\alpha}}\leq \left\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f_2(X_i)\right\|_{\psi_{\alpha}}+\left\|\mathbb{E}\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f_2(X_i)\right\|_{\psi_{\alpha}}\leq 2\left\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f_2(X_i)\right\|_{\psi_{\alpha}}\leq 2\left\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f(X_i)\right\|_{\psi_{\alpha}},
$$

where the above inequality holds because  $\|\cdot\|_{\psi_{\alpha}}$  ( $\alpha = 1, 2$ ) is a standard norm. Then, by Theorem 6.21 of [\[25\]](#page-10-17), we obtain

861  
\n862  
\n863  
\n
$$
\left\| \sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m (f_2(X_i) - \mathbb{E} f_2(X_i)) \right\|_{\psi_\alpha} \leq K_\alpha \left\| \max_{1 \leq i \leq m} \sup_{f \in \mathcal{F}} f(X_i) \right\|_{\psi_\alpha},
$$

which implies

$$
\mathbb{P}\left(\sup_{f_2 \in \mathcal{F}_2} \sum_{i=1}^m f_2(X_i) - \mathbb{E} f_2(X_i) \ge \varepsilon t\right) \le 2 \exp\left(-\left(\frac{\varepsilon t}{K \left\|\max_{1 \le i \le n} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_\alpha}}\right)^2\right). \tag{9}
$$

Next, let us choose  $\varepsilon < 1/10$  and such that

$$
(1 - 5\varepsilon)^{-2} (1 + \beta/2) \le (1 + \beta). \tag{10}
$$

<span id="page-16-1"></span><span id="page-16-0"></span>.

Since  $\varepsilon$  is a function of  $\beta$ , in view of [\(7\)](#page-15-1) and [\(8\)](#page-15-2), we can choose the constant  $K(\alpha, \eta, \beta)$  in [\(1\)](#page-14-4) to be large enough to assure that

$$
B \le 8 \mathbb{E} \max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i) \le \varepsilon t.
$$

**879** Notice that for every  $f \in \mathcal{F}$ , we have  $\mathbb{E}(f_1(X_i) - \mathbb{E}f_1(X_i))^2 \leq \mathbb{E}f_1(X_i)^2 \leq \mathbb{E}f(X_i)^2$ .

Thus, using inequalities [\(5\)](#page-14-3), [\(6\)](#page-15-0), [\(9\)](#page-16-0) and Theorem [8](#page-12-0) (applied for  $\eta$  and  $\beta/2$ ), we obtain

$$
\mathbb{P}(Q - (1 + \eta)\mathbb{E}Q \ge t), \quad \mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t)
$$
  

$$
\le \exp\left(-\frac{t^2(1 - 5\varepsilon)^2}{2(1 + \beta/2)m\sigma^2}\right) + \exp\left(-\frac{(1 - 5\varepsilon)t}{K(\alpha, \eta, \beta)\rho}\right)
$$
  

$$
+ 2\exp\left(-\left(\frac{\varepsilon t}{K_{\alpha} \|\max_{1 \le i \le m} \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_{\alpha}}}\right)^{\alpha}\right)
$$

Since  $\varepsilon$  < 1/10, using [\(7\)](#page-15-1) we can see that for all t with  $K(\alpha, \eta, \beta)$  large enough, we have

$$
\exp\left(-\frac{(1-5\varepsilon)t}{K(\alpha,\eta,\beta)\rho}\right), \exp\left(-\left(\frac{\varepsilon t}{K_{\alpha}\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f(X_{i})\|_{\psi_{\alpha}}}\right)^{\alpha}\right)
$$

$$
\leq \exp\left(-\left(\frac{t}{\widetilde{C}(\alpha,\eta,\beta)\|\max_{1\leq i\leq m}\sup_{f\in\mathcal{F}}f(X_i)\|_{\psi_\alpha}}\right)^{\alpha}\right).
$$

Therefore, for all t,

$$
\mathbb{P}(Q - (1 + \eta)\mathbb{E}Q \ge t), \quad \mathbb{P}(Q - (1 - \eta)\mathbb{E}Q \le -t)
$$
  
\$\le \exp\left(-\frac{t^2(1 - 5\varepsilon)^2}{2(1 + \beta/2)m\sigma^2}\right) + 3\exp\left(-\left(\frac{t}{\widetilde{C}(\alpha, \eta, \beta)\|\max\_{1 \le i \le m} \sup\_{f \in \mathcal{F}} f(X\_i)\|\_{\psi\_\alpha}\right)^\alpha\right).

Finally, we use [\(10\)](#page-16-1) to finish the proof.

 $\Box$ 

 $\setminus^2$ .

<span id="page-16-2"></span>Lemma 5. (Moment-generating function inequality for suprema of empirical process corresponding to classes of sub-Gaussian functions) Let  $X$  and  $Q$  be defined in Theorem [9,](#page-13-0) then for all  $0 < \eta < 1$  and  $\beta > 0$ , there exists a constant  $C = C(\eta, \beta)$ , such that

$$
\mathbb{E}\exp(\lambda(Q-(1+\eta)\mathbb{E}Q)) \le \exp\left(4(1+\beta)m\sigma^2\lambda^2\right) + 3\exp\left(2\left(C\lambda\left\|\max_{i}\sup_{f\in\mathcal{F}}f(X_i)\right\|_{\psi_2}\right)^2\right).
$$

<span id="page-16-3"></span>**913 914 915** Lemma 6. (Moment-generating function inequality for suprema of empirical process corresponding to classes of sub-exponential functions) Let  $X$  and  $Q$  be defined in Theorem [10,](#page-13-1) then for all  $0 < \eta < 1$  and  $\beta > 0$ , there exists a constant  $C = C(\eta, \beta)$ , such that

 $\mathbb{E} \exp(\lambda (Q - (1+\eta) \mathbb{E} Q)) \leq \exp\left(4(1+\beta)m\sigma^2\lambda^2\right) + \exp\left(12\left(C\lambda\left\|\max_{i}\sup_{f\in\mathcal{F}}f(X_i)\right\|_{\psi_1}\right) \right)$ 

$$
917\\
$$

**918 919 920 921** *Proof of Lemma* [5.](#page-16-2) In the proof we use the notation  $\leq$  between two positive sequences  $(a_k)_k$  and  $(b_k)_k$ , writing  $a_k \leq b_k$ , if there exists a constant  $C > 0$  such that for all integer  $k, a_k \leq Cb_k$ . According to Theorem [9,](#page-13-0) we have

$$
\mathbb{P}(|Q - (1 + \eta)\mathbb{E}Q| \ge t) \le 2\exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + 6\exp\left(-\frac{t^2}{C^2\|\max_i\sup_{f \in \mathcal{F}}|f(X_i)\|\|_{\psi_2}^2}\right)
$$

.

Let the random variable  $Y = Q - (1 + \eta) \mathbb{E}Q$  we have that for any  $k \ge 1$ ,

$$
\mathbb{E}[|Y|^k]
$$
  
\n
$$
= \int_0^\infty \mathbb{P}(|Y|^k > t) dt
$$
  
\n
$$
\leq \int_0^\infty 2 \exp\left(-\frac{t^{2/k}}{2(1+\beta)m\sigma^2}\right) dt + \int_0^\infty 6 \exp\left(-\frac{t^{2/k}}{C^2 \|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_2}^2}\right) dt
$$
  
\n
$$
= (2(1+\beta)m\sigma^2)^{k/2} k \int_0^\infty e^{-u} u^{k/2-1} du + 3k \left(C \left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}^2\right)^k \int_0^\infty e^{-v} v^{k/2-1} dv
$$
  
\n
$$
= (2(1+\beta)m\sigma^2)^{k/2} k \Gamma(k/2) + 3k \left(C \left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)^k \Gamma(k/2),
$$

**941 942 943**

> where we denote  $u = \frac{t^{2/k}}{2(1+\beta)m\sigma^2}$  and  $v = \frac{t^{2/k}}{C^2 \|\max_i \sup_{f \in \mathcal{F}_i} u\|_{\infty}}$  $\frac{t^{2}}{\sqrt{C^{2}} \|\max_{i} \sup_{f \in \mathcal{F}} |f(X_{i})| \|_{\psi_{2}}^{2}}$  in the third equality.

Next, we use the Taylor expansion of the exponential function as follows. For  $\lambda > 0$ , we have

$$
\mathbb{E} \exp(\lambda Y)
$$
\n
$$
=1 + \sum_{k=2}^{\infty} \frac{\lambda^{k} \mathbb{E}[|Y|^{k}]}{k!}
$$
\n
$$
\lesssim 1 + \sum_{k=2}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k/2} k \Gamma(k/2) + 3k(C\lambda || \max_{i} \sup_{f \in \mathcal{F}} f(X_{i}) ||_{\psi_{2}})^{k} \Gamma(k/2)}{k!}
$$
\n
$$
=1 + \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k} 2k \Gamma(k)}{(2k)!} + \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k+1/2} (2k+1) \Gamma(k+1/2)}{(2k+1)!}
$$
\n
$$
+ \sum_{k=1}^{\infty} \frac{6k(C\lambda || \max_{i} \sup_{f \in \mathcal{F}} f(X_{i}) ||_{\psi_{2}})^{2k} \Gamma(k)}{k!}
$$
\n
$$
+ \sum_{k=1}^{\infty} \frac{3(2k+1)(C\lambda || \max_{i} \sup_{f \in \mathcal{F}} f(X_{i}) ||_{\psi_{2}})^{2k+1} \Gamma(k+1/2)}{k!}
$$
\n
$$
\leq 1 + \left(2 + \sqrt{2(1+\beta) m \sigma^{2} \lambda^{2}}\right) \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k} k!}{(2k)!}
$$
\n
$$
+ \left(6 + C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}\right) \sum_{k=1}^{\infty} \frac{(C\lambda || \max_{i} \sup_{f \in \mathcal{F}} f(X_{i}) ||_{\psi_{2}})^{2k} k!}{(2k)!},
$$

**969 970**

**971** where the second equality satisfies because of commutative property of positive convergent series. This implies that

 $\leq 1 + \left(1 + \sqrt{\frac{(1+\beta)m\sigma^2\lambda^2}{2}}\right)$ 

2

 $C\lambda \left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}$ 2

**972 973 974**

**975**

**976 977**

**978 979**

$$
\frac{980}{2}
$$

$$
\frac{1}{982}
$$

$$
\frac{1}{98}
$$

$$
\frac{301}{982}
$$

**981**

$$
\frac{981}{100}
$$

$$
\frac{98}{10}
$$

$$
\begin{array}{c} 98 \\ \text{98} \end{array}
$$

$$
\frac{98}{98}
$$

$$
\begin{array}{c} 9 \\ 9 \end{array}
$$

$$
\begin{array}{c}9 \\ 9 \end{array}
$$

$$
\frac{98}{10}
$$

$$
\begin{array}{c} 981 \\ 982 \end{array}
$$

$$
f_{\rm{max}}
$$

$$
\begin{array}{c} 9 \\ 0 \end{array}
$$

$$
\frac{1}{\sqrt{2}}
$$

$$
\overline{a}
$$

$$
\overline{a}
$$

$$
\mathcal{L}_{\mathcal{A}}
$$

$$
\overline{a}
$$

$$
f_{\rm{max}}
$$

$$
\frac{1}{2}
$$

**983**

$$
f_{\rm{max}}
$$

$$
= \exp\left(2(1+\beta)m\sigma^2\lambda^2\right) +
$$

$$
+ \frac{C\lambda \left|\max_i \sup_{f \in \mathcal{F}} f\left(\lambda\right)\right|}{2}
$$

 $\mathbb{E}[|Y|^k]$ 

 $\mathbb{P}(|Y|^k > t) dt$ 

 $\mathbb{P}\left(|Y|>t^{1/k}\right)dt$ 

 $2 \exp \left( - \frac{t^{2/k}}{2(1 + \alpha)^2} \right)$ 

 $2(1+\beta)m\sigma^2$ 

 $=$  $\int^{\infty}$ 0

 $=$  $\int^{\infty}$ 0

 $\leq \int_{0}^{\infty}$ 0

 $+$  $\sqrt{ }$ 3 +

 $\mathbb{E} \exp(\lambda Y)$ 

$$
\frac{C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}}{2} \left(\exp\left(\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)\right)\right)
$$

 $\sum_{\infty}$  $k=1$   $(2(1+\beta)m\sigma^2\lambda^2)^k$  $(2k)!$ 

> $(C\lambda\|\max_i \sup_{f\in\mathcal{F}}f(X_i)\|_{\psi_2})^{2k}$  $(2k)!$

> > $\bigg\|_{\psi_2}$  $\setminus$  $\overline{1}$

 $\left\langle \right\rangle$ 

 $\vert \cdot$ 

 $\left\langle \right\rangle$ 

 $\Big\}$ 

 $\setminus$ 

 $\Big\}$ 

 $\frac{\beta m \sigma^2 \lambda^2}{2}$  (exp  $\left(2(1+\beta)m\sigma^2\lambda^2\right)-1$ )

 $\sum_{\infty}$  $k=1$ 

$$
+3\exp\left(\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)^2\right)
$$
  

$$
\leq \exp\left(4(1+\beta)m\sigma^2\lambda^2\right) + 3\exp\left(2\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)\right)
$$

 $\sqrt{(1+\beta)m\sigma^2\lambda^2}$ 

**993 994 995**

where the first inequality follows from the inequality that  $2(k!)^2 \leq (2k)!$ . The proof is complete.



**1002 1003** *Proof of Lemma [6.](#page-16-3)* According to Theorem [10,](#page-13-1) we have

$$
\mathbb{P}(|Q - (1 + \eta)\mathbb{E}Q| \ge t) \le 2 \exp\left(-\frac{t^2}{2(1 + \beta)m\sigma^2}\right) + 6 \exp\left(-\frac{t}{C\|\max_i \sup_{f \in \mathcal{F}} |f(X_i)\|_{\psi_1}}\right).
$$

Similarly, let the random variable  $Y = Q - (1 + \eta) \mathbb{E}Q$  we have that for any  $k \ge 1$ ,

 $\bigg) dt + \int_{-\infty}^{\infty}$ 0

$$
\begin{array}{c} 1012 \\ 1013 \\ 1014 \\ 1015 \\ 1016 \\ 1017 \\ 1018 \\ 1019 \\ 1020 \\ 1021 \\ 1022 \\ 1023 \end{array}
$$

**1024 1025**

$$
= (2(1+\beta)m\sigma^2)^{k/2} k \int_0^\infty e^{-u} u^{k/2-1} du + 6k \left(C \left\| \max_i \sup_{f \in \mathcal{F}} f(X_i) \right\| \right)
$$
  

$$
\le (2(1+\beta)m\sigma^2)^{k/2} k \Gamma(k/2) + 6k \left(C \left\| \max_i \sup_{f \in \mathcal{F}} f(X_i) \right\|_{\psi_1} \right)^k \Gamma(k),
$$

where we denote 
$$
u = \frac{t^{2/k}}{2(1+\beta)m\sigma^2}
$$
 and  $v = \frac{t^{1/k}}{C \|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_1}}$  in the third equality.

 $6 \exp \left(-\frac{t^{1/k}}{\sqrt{t}}\right)$ 

 $\sqrt{ }$ 

 $C\|\max_i \sup_{f\in\mathcal{F}}f(X_i)\|_{\psi_1}$ 

 $\bigg\|_{\psi_1}$  $\setminus$  $\overline{1}$   $\bigg)$ dt

 $e^{-v}v^{k-1}dv$ 

 $\int^k$ 

**1067**

**1077**

**1026**

$$
\frac{1}{2C\left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}}, \text{ we have}
$$
\n
$$
\mathbb{E} \exp(\lambda Y)
$$
\n
$$
=1 + \sum_{k=2}^{\infty} \frac{\lambda^{k} \mathbb{E}[|Y|^{k}]}{k!}
$$
\n
$$
\leq 1 + \sum_{k=2}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k/2} k \Gamma(k/2) + 6k(C\lambda || \max_{i} \sup_{f \in \mathcal{F}} f(X_{i})||_{\psi_{1}})^{k} \Gamma(k)}{k!}
$$
\n
$$
=1 + \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k} 2k \Gamma(k)}{(2k)!} + \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k+1/2} (2k+1) \Gamma(k+1/2)}{(2k+1)!}
$$
\n
$$
+ \sum_{k=2}^{\infty} 6 \left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}\right)^{k}
$$
\n
$$
\leq 1 + \left(2 + \sqrt{2(1+\beta) m \sigma^{2} \lambda^{2}}\right) \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k} k!}{(2k)!}
$$
\n
$$
+ 6 \left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}\right)^{2} \sum_{k=0}^{\infty} \left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}\right)^{k}
$$
\n
$$
\leq 1 + \left(1 + \sqrt{\frac{(1+\beta) m \sigma^{2} \lambda^{2}}{2}}\right) \sum_{k=1}^{\infty} \frac{(2(1+\beta) m \sigma^{2} \lambda^{2})^{k}}{(2k)!} + 12 \left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}\right)^{2}
$$
\n
$$
=
$$

Next, we use the Taylor expansion of the exponential function as follows. For  $0 \leq \lambda \leq$ 

where the second equality satisfies because of commutative property of positive convergent series and the third inequality follows from the inequality that  $2(k!)^2 \leq (2k)!$  and  $0 \leq \lambda \leq$  $\frac{1}{2C\left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_1}}$ .

**1068** The proof is complete.

 $\Box$ 

## B PROOFS OF SECTION [3](#page-2-0)

**1074 1075 1076** *Proof of Theorem [1.](#page-2-2)* Let  $\{U_1, \ldots, U_m\}$  and  $\{W_1, \ldots, W_m\}$  be sampled uniformly from a finite set of M-dimensional vectors  $^2 \{v_1, \ldots, v_N\} \subset \mathbb{R}^M$  $^2 \{v_1, \ldots, v_N\} \subset \mathbb{R}^M$  $^2 \{v_1, \ldots, v_N\} \subset \mathbb{R}^M$  with and without replacement respectively, where

<span id="page-19-0"></span>**<sup>1078</sup> 1079** <sup>2</sup>We assume that  $F$  is a countable class of functions and this can be translated to the uncountable classes. For instance, if the empirical process is separable, meaning that  $\mathcal F$  contains a dense countable subset. We refer to page 314 of [\[3\]](#page-9-7) or page 72 of [\[5\]](#page-9-1)

1080 
$$
\mathbf{v}_{j} = (f_{1}(c_{j}), \ldots, f_{M}(c_{j}))^{T}.
$$
 According to Lemma 2 and Theorem 6, we get that for all  $\lambda > 0$ :  
\n1082 
$$
\mathbb{E}\left[e^{\lambda Q'_{m}}\right] = \mathbb{E}\left[\exp\left(\lambda \sup_{j=1,\ldots,M}\left(\sum_{i=1}^{m} W_{i}\right)_{j}\right)\right] \leq \mathbb{E}\left[\exp\left(\lambda \sup_{j=1,\ldots,M}\left(\sum_{i=1}^{m} u_{i}\right)_{j}\right)\right] = \mathbb{E}\left[e^{\lambda Q_{m}}\right],
$$
\n1084 (11)

where the lower index  $j$  indicates the  $j$ -th coordinate of a vector. According to Lemma [5,](#page-16-2) the moment generalization function of  $Q_m$  can be bounded, which we can derive the following inequalities

$$
\mathbb{E}\left[e^{\lambda Q'_m}\right] \leq \mathbb{E}\left[e^{\lambda Q_m}\right] \leq \exp\left((1+\eta)\lambda \mathbb{E}[Q_m] + 4(1+\beta)m\sigma^2\lambda^2\right) + 3\exp\left((1+\eta)\lambda \mathbb{E}[Q_m] + 2\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)^2\right)
$$

**1093** or, equivalently,

<span id="page-20-1"></span>**1109**

$$
\mathbb{E}\left[e^{\lambda(Q'_m - (1+\eta)\mathbb{E}[Q'_m])}\right]
$$
  
\n
$$
\leq \exp\left((1+\eta)\lambda(\mathbb{E}[Q_m] - \mathbb{E}[Q'_m]) + 4(1+\beta)m\sigma^2\lambda^2\right)
$$
  
\n
$$
+ 3\exp\left((1+\eta)\lambda(\mathbb{E}[Q_m] - \mathbb{E}[Q'_m]) + 2\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}\right)^2\right).
$$

**1102** Using Chernoff's method, we can obtain that for all  $\epsilon \geq 0$  and  $\lambda > 0$ :

1103 
$$
\mathbb{P}\left\{Q'_{m} - (1+\eta)\mathbb{E}[Q'_{m}]\right\} \geq \epsilon
$$
\n1104 
$$
\mathbb{E}\left[e^{\lambda(Q'_{m} - (1+\eta)\mathbb{E}[Q'_{m}])}\right]
$$
\n1105 
$$
\leq \frac{\mathbb{E}\left[e^{\lambda(Q'_{m} - (1+\eta)\mathbb{E}[Q'_{m}])}\right]}{e^{\lambda\epsilon}}
$$
\n1106 
$$
\leq \exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) + 4(1+\beta)m\sigma^{2}\lambda^{2}\right)}
$$
\n1107 
$$
+ \frac{3\exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) + 2\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}\right)^{2}\right)}{\exp(\lambda\epsilon)}
$$
\n1111 
$$
+ \frac{1112}{\exp((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]))\left(\exp(4(1+\beta)m\sigma^{2}\lambda^{2}) + 3\exp\left(2\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}\right)^{2}\right)\right)}{\exp(\lambda\epsilon)}
$$
\n1115 
$$
\leq 6\exp\left(((1+\eta)(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) - \epsilon)\lambda + \left(4(1+\beta)m\sigma^{2} + 2C^{2} \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}^{2}\right)\lambda^{2}\right),
$$
\n1117 
$$
1118
$$
 where the first inequality applies Chernoff's method. The third hold under the following two terms

**1119 1120 1121 1122** where the first inequality applies Chernoff's method. The third hold under the following two terms  $\exp\left( 4(1+\beta)m\sigma^2\lambda^2\right)\,\geq\,1$  and  $\exp\left( 2\left( C\lambda\left\|\max_i\sup_{f\in\mathcal{F}}f(X_i)\right\|_{\psi_2}\right)$  $\binom{2}{1} \geq 1$ . Using  $a + b \leq$ 2ab,  $\forall a, b \geq 1$ , we obtain the third inequality.

**1123** The term on the right-hand side of the last inequality achieves its minimum for

<span id="page-20-0"></span>
$$
\lambda = \frac{\epsilon + (1+\eta)\left(\mathbb{E}[Q_m'] - \mathbb{E}[Q_m]\right)}{8(1+\beta)m\sigma^2 + 4C^2 \left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}^2}.
$$
\n(13)

**1128 1129 1130 1131** Insert [\(13\)](#page-20-0) into [\(12\)](#page-20-1), when we have the technical condition  $\epsilon \geq (1 + \eta)(\mathbb{E}[Q_m] - \mathbb{E}[Q'_m])$  where  $\mathbb{E}[Q_m] \geq \mathbb{E}[Q'_m]$  follows from Theorem [6](#page-12-2) by exploiting the fact that the supremum is a convex function., we obtain the following inequality

1131  
\n
$$
\mathbb{P}\left\{Q'_m - (1+\eta)\mathbb{E}[Q_m] \ge \epsilon\right\} \le 6 \exp\left(-\frac{\epsilon^2}{16(1+\beta)m\sigma^2 + 8C^2 \left\|\max_i \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_2}^2}\right).
$$

 $\mathbb{E}\left[e^{\lambda(Q'_m-(1+\eta)\mathbb{E}[Q'_m])}\right]$ 

### **1134 1135** The proof is complete.

**1136**

**1137 1138 1139 1140 1141 1142** *Proof of Theorem [2.](#page-3-0)* The proof of Theorem [2](#page-3-0) is similar with Theorem [1.](#page-2-2) Let two series of random variables  $\{U_1, \ldots, U_m\}$  and  $\{W_1, \ldots, W_m\}$  be sampled uniformly form a finite set of Mdimensional vectors  $\{v_1, \ldots, v_N\}$  ⊂  $\mathbb{R}^M$  with and without replacement respectively, where  $\mathbf{v}_j = (f_1(c_j), \dots, f_M(c_j))^T$ . According to Lemma [2](#page-12-1) and Theorem [6,](#page-12-2) we get that for all  $\lambda \leq 0$ :  $\lceil$  /  $\setminus$  ) ]  $\lceil$  /  $\setminus$  ) ]

 $\Box$ 

$$
\mathbb{E}\left[e^{\lambda Q'_m}\right] = \mathbb{E}\left[\exp\left(\lambda \sup_{j=1,\dots,M}\left(\sum_{i=1}^m W_i\right)_j\right)\right] \leq \mathbb{E}\left[\exp\left(\lambda \sup_{j=1,\dots,M}\left(\sum_{i=1}^m u_i\right)_j\right)\right] = \mathbb{E}\left[e^{\lambda Q_m}\right],\tag{14}
$$

where the lower index  $j$  indicates the  $j$ -th coordinate of a vector. According to Lemma [6,](#page-16-3) the moment generalization function of  $Q_m$  can be bounded, which we can derive the following inequalities

$$
\mathbb{E}\left[e^{\lambda Q'_m}\right] \leq \mathbb{E}\left[e^{\lambda Q_m}\right] \leq \exp\left((1+\eta)\lambda \mathbb{E}[Q_m] + 4(1+\beta)m\sigma^2\lambda^2\right) \n+ \exp\left((1+\eta)\lambda \mathbb{E}[Q_m] + 12\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_i)\right\|_{\psi_1}\right)^2\right)
$$

**1153** or, equivalently,

**1154 1155**

$$
\frac{1156}{1157}
$$

**1158 1159**

**1160 1161**

 $\leq$ exp $((1 + \eta)\lambda(\mathbb{E}[Q_m] - \mathbb{E}[Q_m']) + 4(1 + \beta)m\sigma^2\lambda^2)$  $+$  exp  $\sqrt{ }$  $(1 + \eta)\lambda(\mathbb{E}[Q_m] - \mathbb{E}[Q'_m]) + 12$  $\sqrt{ }$  $\bigcap C$ λ  $\max_i \sup_{f \in \mathcal{F}} f(X_i)$  $\bigg\|_{\psi_1}$  $\setminus$  $\perp$  $^{2}$  $\vert \cdot$ 

**1162 1163** Using Chernoff's method, we can obtain that for all  $\epsilon \ge 0$  and  $0 \le \lambda \le \frac{1}{2C \|\max_i \sup_{f \in \mathcal{F}} f(X_i)\|_{\psi_1}}$ :

<span id="page-21-1"></span>1164 
$$
\mathbb{P}\left\{Q'_{m} - (1+\eta)\mathbb{E}[Q'_{m}]\right\} \geq \epsilon\}
$$
\n1165 
$$
\leq \frac{\mathbb{E}\left[e^{\lambda(Q'_{m} - (1+\eta)\mathbb{E}[Q'_{m}])}\right]}{e^{\lambda\epsilon}}
$$
\n1166 
$$
\leq \frac{\exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) + 4(1+\beta)m\sigma^{2}\lambda^{2}\right)}{\exp(\lambda\epsilon)}
$$
\n1168 
$$
\leq \frac{\exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) + 12\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}\right)^{2}\right)}{\exp(\lambda\epsilon)}
$$
\n1171 
$$
+ \frac{\exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}])\right)\left(\exp(4(1+\beta)m\sigma^{2}\lambda^{2}) + \exp\left(12\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}\right)^{2}\right)\right)}{\exp(\lambda\epsilon)}
$$
\n1175 
$$
\leq \frac{\exp\left((1+\eta)\lambda(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}])\right)\left(\exp(4(1+\beta)m\sigma^{2}\lambda^{2}) + \exp\left(12\left(C\lambda \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{2}}\right)^{2}\right)\right)}{\exp(\lambda\epsilon)}
$$
\n1176 
$$
\leq 2\exp\left(((1+\eta)(\mathbb{E}[Q_{m}] - \mathbb{E}[Q'_{m}]) - \epsilon)\lambda + \left(4(1+\beta)m\sigma^{2} + 12C^{2}\left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}^{2}\right)\lambda^{2}\right), \tag{15}
$$

**1179 1180 1181 1182 1183** where the first inequality applies Chernoff's method and the third hold under the following two terms  $\exp\left(4(1+\beta)m\sigma^2\lambda^2\right)\geq 1$  and  $\exp\left(2\left(C\lambda\left\|\max_i\sup_{f\in\mathcal{F}}f(X_i)\right\|_{\psi_2}\right)$  $\binom{2}{1} \geq 1$ . Using  $a +$  $b \le 2ab$ ,  $\forall a, b \ge 1$ , we obtain the third inequality.

**1184** The term on the right-hand side of the last inequality achieves its minimum for

<span id="page-21-0"></span>
$$
\lambda = \frac{\epsilon + (1+\eta)(\mathbb{E}[Q_m'] - \mathbb{E}[Q_m])}{\epsilon}
$$

$$
\lambda = \frac{c + (1 + \eta)\ln|\mathcal{L}(\mathcal{L}_{m})|}{8(1 + \beta)m\sigma^{2} + 24C^{2} \left\|\max_{i} \sup_{f \in \mathcal{F}} f(X_{i})\right\|_{\psi_{1}}^{2}}.
$$
(16)

**1188 1189 1190** Insert [\(16\)](#page-21-0) into [\(15\)](#page-21-1), when we have the technical condition  $(1 + \eta)(\mathbb{E}[Q_m] - \mathbb{E}[Q'_m]) \le \epsilon$  $12C \left\| \max_i \sup_{f \in \mathcal{F}} f(X_i) \right\|_{\psi_1}$ , we obtain the following inequality

1191  
\n
$$
\mathbb{P}\left\{Q'_m - (1+\eta)\mathbb{E}[Q_m] \ge \epsilon\right\} \le 2\exp\left(-\frac{\epsilon^2}{16(1+\beta)m\sigma^2 + 48C^2\left\|\max_i \sup_{f \in \mathcal{F}}f(X_i)\right\|_{\psi_1}^2}\right).
$$
\n1194

The proof is complete.

**1196 1197**

**1191**

**1194 1195**

**1198 1199 1200**

## C PROOFS OF SECTION [4](#page-3-1)

**1201 1202** C.1 PROOFS OF SUBSECTION [4.1](#page-4-0)

**1203 1204** From now on it will be convenient to introduce the following operators, mapping functions  $f$  defined on  $X_N$  to  $\mathbb{R}$ :

$$
Ef = \frac{1}{N} \sum_{i=1}^{N} f(\mathbf{x}_i), \mathbf{x}_i \in \mathbf{X}_N, \quad E_m f = \frac{1}{N} \sum_{\mathbf{x}_j=1}^{m} f(\mathbf{x}_j), \mathbf{x}_j \in \mathbf{X}_m.
$$

**1209 1210 1211** Assume that there is a function  $\mathbf{w}_N^* \in \mathcal{W}$  satisfying  $R_N(\mathbf{w}_N^*) = \inf_{\mathbf{w} \in \mathcal{W}} R_N(\mathbf{w})$ . Define the excess loss class  $\mathcal{F}^* = \{f : f(\mathbf{x}) = \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}_N^*; \mathbf{x}), \mathbf{w} \in \mathcal{W}\}.$ 

**1212 1213** Let  $\{\xi_1, \ldots, \xi_n\}$  be random variables sampled with replacement from  $X_N$ . The mapping functions f defined on  $X_N$  to  $\mathbb R$ . Denote

$$
E_{r,m}f = \mathbb{E}\left[\sup_{f \in \mathcal{F}^*:Ef^2 \le r} \left(Ef - \frac{1}{m} \sum_{i=1}^m f(\xi_i)\right)\right].
$$
 (17)

<span id="page-22-0"></span>1

 $\Box$ 

**1218** Then we have

$$
E_{r,m}f = \mathbb{E}\left[\sup_{f \in \mathcal{F}^*:Ef^2 \le r} \left(Ef - \frac{1}{m} \sum_{i=1}^m f(\xi_i)\right)\right]
$$
  
\n
$$
\le 2\mathbb{E}_{\xi \sim \mathbf{X}_N, v} \left[\sup_{f \in \mathcal{F}^*:Ef^2 \le r} v_i \left(Ef - \frac{1}{m} \sum_{i=1}^m f(\xi_i)\right)\right]
$$
  
\n
$$
\le 2\mathbb{E}_v \left[\sup_{f \in \mathcal{F}^*:Ef^2 \le r} \sum_{i=1}^m v_i Ef\right] + 2\mathbb{E}_{\xi \sim \mathbf{X}_N, v} \left[\sup_{f \in \mathcal{F}^*:Ef^2 \le r} \frac{1}{m} \sum_{i=1}^m v_i f(\xi_i)\right]
$$
  
\n
$$
= 2\Re_N \{f \in \mathcal{F}^*:Ef^2 \le r\}.
$$

**1219 1220**

$$
\begin{array}{c} 1226 \\ 1227 \\ 1228 \end{array}
$$

**1229**

**1235**

**1230** where the first inequality holds using symmetrization inequality (see Lemma 11.4 [\[3\]](#page-9-7))

<span id="page-22-1"></span>**1231 1232 1233 1234** Lemma 7 (Peeling Lemma for sub-Gaussian). *Assume that there is a constant* B > 0 *such that for every*  $f \in \mathcal{F}^*$  *we have*  $Ef^2 \leq B \cdot Ef$ . Suppose Assumptions *1* hold and the objective function  $\ell(\cdot;\cdot)$ *is sub-Gaussian.. Assume there is a sub-root function*  $\psi_m(r)$  *such that* 

$$
2B\Re_N\{f \in \mathcal{F}^* : Ef^2 \le r\} \le \psi_m(r),
$$

**1236 1237** *where*  $E_{r,m}$  was defined in [\(17\)](#page-22-0). Let  $r_m^*$  be a fixed point of  $\psi_m(r)$ .

**1238 1239** *Fix some*  $\lambda > 1$ *. For*  $w(r, f) = \min\{r\lambda^k : k \in \mathbb{N}, r\lambda^k \ge Ef^2\}$ *, define the following rescaled version of excess loss class:*

$$
\mathcal{G}_r = \left\{ \frac{r}{w(r,f)} f : f \in \mathcal{F}^* \right\}.
$$

**1242 1243** *Then for any*  $r > r_m^*$  *and*  $t > 0$ *, with probability at least*  $1 - \delta$ *, we have* 

**1244 1245**

$$
\sup_{g \in \mathcal{G}_r} Eg - E_m g \le \frac{(1+\eta)\sqrt{rr_m^*}}{B} \left(1 + \frac{1}{K_2\sqrt{\log\frac{2}{\delta}}}\right)
$$

**1246 1247 1248**

**1249 1250**

**1252**

$$
+\,4\sqrt{(1+\beta)\left(\frac{N}{m^2}\right)r\log\frac{12}{\delta}}+4\sqrt{\frac{2C^2K\log N}{m^2}\log\frac{12}{\delta}},
$$

**1251** *where*  $K, K_2, \eta, \beta$  *are some positive constants. C is positive constants depending on*  $\eta, \beta$ *.* 

**1253 1254** *Proof of Lemma [7.](#page-22-1)* We use traditional peeling technologies presented in the proof of the first part of Theorem 3.3 of [\[2\]](#page-9-8), but using Theorem [1](#page-2-2) in place of Talagrand's inequality.

**1255 1256** Firstly, for any  $f \in \mathcal{F}^*$ , we have

**1257**

**1258**

<span id="page-23-0"></span> $\mathbb{V}[f(\mathbf{x})] = Ef^2 - (Ef)^2 \le Ef^2$ . (18)

**1259** Let us fix some  $\lambda > 1$  and  $r > 0$  and introduce the following rescaled version of excess loss class:

$$
\mathcal{G}_r = \left\{ \frac{r}{w(r,f)} f : f \in \mathcal{F}^* \right\},\
$$

**1264** where  $w(r, f) = \min\{r\lambda^k : k \in \mathbb{N}, r\lambda^k \ge Ef^2\}.$ 

**1265 1266 1267 1268 1269** Let us consider functions  $f \in \mathcal{F}^*$  such that  $Ef^2 < r$ , meaning  $w(r, f) = r$ . The functions  $g \in \mathcal{G}_r$  corresponding to those functions satisfy  $g = f$  and thus  $\mathbb{V}[g(\mathbf{x})] = \mathbb{V}[f(\mathbf{x})] \leq Ef^2 \leq r$ . Otherwise, if  $Ef^2 > r$ , then  $w(r, f) = \lambda^k r$ , and thus the functions  $g \in \mathcal{G}_r$  corresponding to them satisfy  $g = \frac{f}{\lambda^k}$  and  $Ef^2 \in (r\lambda^{k-1}, r\lambda^k]$ . Thus we have  $\mathbb{V}[g(\mathbf{x})] = \frac{\mathbb{V}[f(\mathbf{x})]}{\lambda^{2k}} \le \frac{Ef^2}{\lambda^{2k}} \le r$ . We conclude that, for any  $g \in \mathcal{G}_r$ , it holds  $\mathbb{V}[g(X)] \leq r$ .

**1270 1271** Next we need to upper bound the following quantity:

$$
V_r = \sup_{g \in \mathcal{G}_r} Eg - E_m g.
$$

**1274 1275** Note that any  $f \in \mathcal{F}^*$ ,  $f(\mathbf{x})$  is sub-Gaussian, thus for all  $g \in \mathcal{G}_r$ ,  $g(\mathbf{x})$  is sub-Gaussian. Notice that

$$
\frac{1}{2}(Eg - E_m g) = \frac{1}{m} \sum_{\mathbf{x} \in \mathbf{X}_m} \frac{Eg - g(\mathbf{x})}{2}.
$$

**1277 1278 1279**

**1276**

**1272 1273**

**1280 1281** Note that  $(Eg - g(\mathbf{x}))/2$  is also sub-Gaussian and  $\mathbb{E}[Eg - g(\mathbf{x})] = 0$ . Since Eg is not random, using [\(18\)](#page-23-0), for all  $g \in \mathcal{G}_r$  we also have

$$
\mathbb{V}\left[\frac{Eg-g(\mathbf{x})}{2}\right]=\frac{\mathbb{V}[g(\mathbf{x})]}{4}\leq \frac{r}{4},
$$

**1284 1285**

**1282 1283**

**1286 1287** Besides, we need to bound  $\left\|\max_{\mathbf{x}} \sup_{g \in \mathcal{G}_r} \frac{Eg - g(\mathbf{x})}{2}\right\|$  $\frac{-g(\mathbf{x})}{2}$ 2  $\ddot{\psi}_2$ .

$$
\left\|\max_{\mathbf{x}} \sup_{g \in \mathcal{G}_r} \frac{E_g - g(\mathbf{x})}{2}\right\|_{\psi_2}^2 = \frac{\left\|\max_{\mathbf{x}} \sup_f Ef - f(\mathbf{x})\right\|_{\psi_2}^2}{4\lambda^{2k}}
$$

$$
\leq K^2 \max_{\mathbf{x}} \left\|\sup_f \ell(\mathbf{w}; \mathbf{x})\right\|_{\psi_2}^2 \log N \leq K \log N,
$$

**1295** where  $K$  is a positive constant. The first inequality holds using Theorem [\[34\]](#page-10-15) and the second inequality satisfies because  $\ell(\cdot; \mathbf{x})$  is sub-Gaussian.

**1296 1297 1298** We can now apply either Theorem [1](#page-2-2) for the following function class:  $\{(Eg - g(\mathbf{x}))/2, g \in \mathcal{G}_r\}.$ Here we present the proof based on Theorem [1.](#page-2-2) Applying it we get that for all  $\delta \in (0,1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we have

**1299 1300 1301**

$$
\frac{1}{2}\sup_{g\in\mathcal{G}_r}Eg - E_m g
$$

2

$$
\leq \frac{1+\eta}{2} \mathbb{E}\left[\sup_{g \in \mathcal{G}_r} E_{r,m}g\right] + \sqrt{\left(16(1+\beta)\left(\frac{N}{m^2}\right)\frac{1}{4}\sup_{g \in \mathcal{G}_r} \mathbb{V}[g(\mathbf{x})] + \frac{8C^2K\log N}{m^2}\right)\log\frac{12}{\delta}}\n\n\leq \frac{1+\eta}{2} \mathbb{E}\left[\sup_{g \in \mathcal{G}_r} E_{r,m}g\right] + \sqrt{\left(4(1+\beta)\left(\frac{N}{m^2}\right)r + \frac{8C^2K\log N}{m^2}\right)\log\frac{12}{\delta}}\n\n\leq \frac{1+\eta}{2} \mathbb{E}\left[\sup_{g \in \mathcal{G}_r} E_{r,m}g\right] + 2\sqrt{(1+\beta)\left(\frac{N}{m^2}\right)r\log\frac{12}{\delta}} + 2\sqrt{\frac{2C^2K\log N}{m^2}\log\frac{12}{\delta}},
$$

<span id="page-24-0"></span> $m<sup>2</sup>$ 

**1308 1309 1310**

**1311**

**1314 1315 1316** where the last inequality holds because  $\sqrt{a+b} \leq \sqrt{a}$  + √ b for any  $a \ge 0$  and  $b \ge 0$ .

**1312 1313** Rewriting above inequality we have

$$
Vr \le (1+\eta)\mathbb{E}\left[\sup_{g\in\mathcal{G}_r} E_{r,m}g\right] + 4\sqrt{(1+\beta)\left(\frac{N}{m^2}\right)r\log\frac{12}{\delta}} + 4\sqrt{\frac{2C^2K\log N}{m^2}\log\frac{12}{\delta}}.\tag{19}
$$

**1317 1318 1319 1320 1321** Now we set  $\mathcal{F}^*(x, y) = \{f \in \mathcal{F}^* : x \leq Ef^2 \leq y\}$ , Note that  $Ef$  is sub-Gaussian, for  $f \in \mathcal{F}^*$ , for any  $\delta \in (0, 1)$  with probability at least  $1 - \frac{\delta}{2}$ , we have  $\mathbb{V}[f(\mathbf{x})] \leq E f^2 \leq B \cdot E f \leq BK_2 \sqrt{\log 2/\delta}$ . Let  $\sum_{i=1}^{n}$  be the smallest integer such that  $r\lambda^{k+1} \leq BK_2\sqrt{\log 2/\delta}$ . Notice that, for any sets A and  $B$ , we have:

$$
\mathbb{E}\left[\sup_{g\in A\cup B} E_{r,m}g\right] \leq \mathbb{E}\left[\sup_{g\in A} E_{r,m}g\right] + \mathbb{E}\left[\sup_{g\in B} E_{r,m}g\right]
$$

**1324 1325 1326** Since supremum is a convex function ,we can use Jensen's inequality to show that each of the terms is positive. Then for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we have:

**1327 1328**

**1322 1323**

$$
\mathbb{E}\left[\sup_{g\in\mathcal{G}_r} E_{r,m}g\right]
$$
\n
$$
\leq \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(0,r)} E_{r,m}f\right] + \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(r,2BK_2\sqrt{2\log 2/\delta})} \frac{r}{w(r,f)} E_{r,m}f\right]
$$
\n
$$
\leq \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(0,r)} E_{r,m}f\right] + \sum_{i=0}^k \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(r\lambda^i,r\lambda^{i+1})} \frac{r}{w(r,f)} E_{r,m}f\right]
$$
\n
$$
\leq \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(0,r)} E_{r,m}f\right] + \sum_{i=0}^k \lambda^{-i} \mathbb{E}\left[\sup_{f\in\mathcal{F}^*(r\lambda^i,r\lambda^{i+1})} E_{r,m}f\right]
$$

$$
\leq 2\Re_N\{f\in\mathcal{F}^*:Ef^2\leq r\}+2\sum_{i=0}^n\lambda^{-i}\Re_N\{f\in\mathcal{F}^*:r\lambda^i\leq Ef^2\leq r\lambda^{i+1}\}
$$

$$
\leq \frac{\psi_m(r)}{B} + \frac{1}{BK_2\sqrt{\log\frac{2}{\delta}}}\sum_{i=0}^k \lambda^{-i}\psi_m(r\lambda^{i+1}),
$$

**1345 1346** where the last inequality satisfies because  $Ef$  is sub-Gaussian. Next, since  $\psi_m$  is sub-root, for any  $\beta \geq 1$ , we have  $\psi_m(\beta r) \leq \sqrt{\beta} \psi_m(r)$ . Thus

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1348  
1349  

$$
\mathbb{E}[V_r] \leq \sqrt{\beta} \leq \frac{\psi_m(r)}{B} \left(1 + \frac{\sqrt{\lambda}}{K_2 \sqrt{\log \frac{2}{\delta}}} \sum_{i=0}^k \lambda^{-i/2}\right).
$$

**1350 1351 1352 1353 1354** Taking  $\lambda = 4$ , the right hand side is upper bounded by  $\frac{\psi_m(r)}{B}$  $\left(1+\frac{1}{K_2\sqrt{\log\frac{2}{\delta}}} \right)$  . Finally we note that for  $r \geq r_m^*$ , then for all  $r \geq r_m^*$ , it holds  $\psi_m(r) \leq \sqrt{r/r_m^*} \psi_m(r_m^*) = \sqrt{rr_m^*}$ . Thus, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$ 

> $\mathbb{E} \left[ \frac{\sqrt{2}}{2} \right]$  $\left[\sup_{g\in\mathcal{G}_r} E_{r,m}g\right]\leq$  $\sqrt{rr_m^*}$ B  $\sqrt{ }$  $\left(1+\frac{1}{\sqrt{1}}\right)$  $K_2\sqrt{1\log\frac{2}{\delta}}$  $\setminus$  $\left| \begin{array}{ccc} \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \end{array} \right|$

**1359 1360** Combining [\(20\)](#page-25-0) and [\(19\)](#page-24-0), according to the union bound, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , we have

$$
\sup_{g \in \mathcal{G}_r} Eg - E_m g \le \frac{(1+\eta)\sqrt{rr_m^*}}{B} \left(1 + \frac{1}{K_2\sqrt{\log\frac{2}{\delta}}}\right) + 4\sqrt{(1+\beta)\left(\frac{N}{m^2}\right)r\log\frac{12}{\delta}} + 4\sqrt{\frac{2C^2K\log N}{m^2}\log\frac{12}{\delta}},
$$

**1368 1369 1370** where  $K, K_2, \eta, \beta$  are some positive constants. C is positive constants depending on  $\eta, \beta$ . The proof is complete.

<span id="page-25-0"></span> $\Box$ 

<span id="page-25-3"></span>,

<span id="page-25-1"></span>**1372 1373 1374 Lemma 8.** *Under the assumptions of Theorem [3,](#page-4-2) for any*  $\delta \in (0,1)$ *, with probability at least*  $1-\delta$ *, we have*

$$
\begin{array}{c} 1375 \\ 1376 \end{array}
$$

**1377**

**1380 1381**

**1371**

$$
R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}^*_N) \leq \frac{c_1r_m^*}{B\log\frac{2}{\delta}} + \frac{c_2N\log\frac{12}{\delta}}{m^2} + \frac{c_3\sqrt{\log N\log\frac{12}{\delta}}}{m}
$$

**1378 1379** *where*  $c_1$ ,  $c_2$  *and*  $c_3$  *are some positive constants.* 

*Proof of Lemma* [8.](#page-25-1) According to Lemma [7,](#page-22-1) we have the following results that, for any  $r > r_m^*$ ,  $\delta \in (0, 1)$  and  $\lambda > 1$ , with probability at least  $1 - \delta$ , we have

$$
\begin{array}{c} 1382 \\ 1383 \\ 1384 \end{array}
$$

**1391 1392 1393**

<span id="page-25-2"></span>
$$
\sup_{g \in \mathcal{G}_r} Eg - E_m g \le \frac{(1+\eta)\sqrt{rr_m^*}}{B} \left(1 + \frac{1}{K_2\sqrt{\log\frac{2}{\delta}}}\right) + 4\sqrt{(1+\beta)\left(\frac{N}{m^2}\right)r\log\frac{12}{\delta}} + 4\sqrt{\frac{2C^2K\log N}{m^2}\log\frac{12}{\delta}},
$$
\n(21)

**1390** where  $\mathcal{G}_r$  is the rescaled excess loss class:

$$
\mathcal{G}_r = \left(\frac{r}{w(r,f)}f : f \in \mathcal{F}^*\right),\,
$$

**1394 1395** and  $w(r, f) = \min\{r\lambda^k : k \in \mathbb{N}, r\lambda^k \ge Ef^2\}$ . Now we want to choose  $r_0 > r_m^*$  in such a way that the upper bound of [\(21\)](#page-25-2) becomes of a form  $\frac{r_0}{\lambda BK'}$ , we achieve this by setting:

$$
r_0 = K'^2 \lambda^2 \left( (1+\eta) \sqrt{r_m^*} \left( 1 + \frac{1}{K_2 \sqrt{\log \frac{2}{\delta}}} \right) + 4B \sqrt{(1+\beta) \left( \frac{N}{m^2} \right) \log \frac{12}{\delta}} \right)^2 > r_m^*.
$$

Inserting  $r = r_0$  into [\(21\)](#page-25-2), we have

1402  
1403  

$$
\sup_{g \in \mathcal{G}_{r_0}} Eg - E_m g \le \frac{r_0}{\lambda B K'} + 4\sqrt{\frac{2C^2 K \log N}{m^2} \log \frac{12}{\delta}}.
$$
 (22)

**1404 1405** Further, using inequality  $(u + v)^2 \le 2(u^2 + v^2)$ , we have

**1406 1407**

**1408 1409**

**1411 1412**

**1415 1416 1417**

$$
r_0 \le 2(1+\eta)^2 \left(1 + \frac{1}{K_2\sqrt{\log\frac{2}{\delta}}}\right)^2 K'^2 \lambda^2 r_m^* + 32(1+\beta) \left(\frac{N}{m^2}\right) K'^2 \lambda^2 B^2 \log\frac{12}{\delta}.\tag{23}
$$

**1410** Recall that for any  $r > 0$  and all  $g \in \mathcal{G}_r$ , the following holds with probability 1

<span id="page-26-0"></span>
$$
Eg - E_m g \le \sup_{g \in \mathcal{G}_r} Eg - E_m g.
$$

**1413 1414** Using the definition of  $\mathcal{G}_r$ , for all  $f \in \mathcal{F}^*$ , with probability 1, we have the following inequality

$$
E\left(\frac{r}{w(r,f)}f\right) - E_m\left(\frac{r}{w(r,f)}f\right) \le \sup_{g \in \mathcal{G}_r} E_g - E_m g,
$$

**1418** or, rewriting

$$
Ef - E_m f \le \frac{w(r, f)}{r} \sup_{g \in \mathcal{G}_r} E_g - E_m g.
$$

**1420 1421 1422**

**1419**

Next we setting  $r = r_0$  and using [\(22\)](#page-25-3), for any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , we have

$$
\forall f \in \mathcal{F}^*, \forall K > 1: \quad Ef - E_m f \le \frac{w(r_0, f)}{r_0} \left( \frac{r_0}{\lambda K' B} + 4 \sqrt{\frac{2C^2 K \log N}{m^2} \log \frac{12}{\delta}} \right).
$$

**1427 1428 1429** Next, according to  $Ef^2 \leq B \cdot Ef$ , if for  $f \in \mathcal{F}^*, Ef^2 \leq r_0$ , we have  $w(r_0, f) = r_0$  and using [\(23\)](#page-26-0), we have

1430  
\n1431  
\n1432  
\n1433  
\n1434  
\n1435  
\n
$$
\frac{2(1+\eta)^2 K' \lambda r_m^*}{B} \left(1 + \frac{1}{K_2 \sqrt{\log \frac{2}{\delta}}} \right)^2 + 32(1+\beta) \left(\frac{N}{m^2}\right) K' \lambda B \log \frac{12}{\delta} + 4\sqrt{\frac{2C^2 K \log N}{m^2} \log \frac{12}{\delta}}.
$$

Rewriting,

<span id="page-26-1"></span>
$$
Ef \le E_m f + \frac{2(1+\eta)^2 K' \lambda r_m^*}{B} \left(1 + \frac{1}{K_2 \sqrt{\log \frac{2}{\delta}}}\right)^2
$$
  
+ 32(1+\beta)  $\left(\frac{N}{m^2}\right) K' \lambda B \log \frac{12}{\delta} + 4 \frac{CK_2 \sqrt{2 \log \frac{12}{\delta}}}{m}.$  (24)

**1441 1442**

**1446 1447**

**1443 1444 1445** On the other hand, if  $Ef^2 > r_0$ , then  $w(r_0, f) = \lambda^i r_0$  for certain value of  $i > 0$  and also  $Ef^2 \in$  $(r_0 \lambda^{i-1}, r_0 \lambda^i]$ . Then we have

$$
Ef - E_m f
$$

1448  
\n1449  
\n
$$
\leq \frac{w(r_0, f)}{r_0} \left( \frac{r_0}{\lambda K'B} + 4\sqrt{\frac{2C^2 K \log N}{m^2} \log \frac{12}{\delta}} \right)
$$

1450  
\n1451  
\n
$$
\leq \frac{\lambda^{i-1}r_0}{K'B} + \frac{4\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{12}{\delta}}}{m}
$$

- K′B
- **1453 1454**  $\leq \frac{Ef^2}{V}$  $4\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{12}{\delta}}$

$$
1455 \t\t\t m
$$

$$
\leq E f \atop 1457 \n\leq \frac{Ef}{K'} + \frac{4\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{12}{\delta}}}{m}.
$$

**1458 1459** Thus, we have

**1460 1461**

**1462**

<span id="page-27-0"></span>
$$
Ef \leq \frac{K'}{K'-1} E_m f + \frac{4K'\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{12}{\delta}}}{(K'-1)m}.
$$
 (25)

**1463** Combing [\(24\)](#page-26-1) and [\(25\)](#page-27-0), for any  $\delta \in (0,1)$ , with probability at least  $1 - \delta$ , we have

$$
\forall f \in \mathcal{F}^*, \forall K > 1: \quad Ef \le \inf_{K' > 1} \frac{K'}{K' - 1} E_m f + \frac{2(1 + \eta)^2 K' \lambda r_m^*}{B} \left( 1 + \frac{1}{K_2 \sqrt{\log \frac{2}{\delta}}} \right)^2 \tag{26}
$$

<span id="page-27-2"></span>
$$
+32(1+\beta)\left(\frac{N}{m^{2}}\right)K^{\prime}\lambda B\log\frac{12}{\delta}+4\frac{CK_{2}\sqrt{2\log\frac{12}{\delta}}}{m}+\frac{4K^{\prime}\lambda^{i-1}\sqrt{2C^{2}K\log N\log\frac{12}{\delta}}}{(K-1)m}.
$$

Finally we recall that the definition of  $\mathcal{F}^*$  and put  $\hat{f}_m(\cdot) = \ell(\hat{\mathbf{w}}_m; \cdot) - \ell(\mathbf{w}_N^*; \cdot)$ . Notice that

$$
E_m\hat{f}_m = E_m\ell(\hat{\mathbf{w}}_m) - E_m\ell(\mathbf{w}_N^*) = \hat{R}_m(\hat{\mathbf{w}}_m) - \hat{R}_m(\mathbf{w}_N^*) \leq 0,
$$

**1475 1476**

**1477 1478**

$$
E\hat{f}_m = R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*),
$$

**1479** thus, we have

and

$$
R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*) \le \frac{c_1 r_m^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{m},
$$

**1484** where  $c_1$ ,  $c_2$  and  $c_3$  are some positive constants.

**1485** The proof is complete.

<span id="page-27-1"></span>**1486 1487 1488 Lemma 9.** *Under the assumptions of Theorem [3,](#page-4-2) for any*  $\delta \in (0,1)$ *, with probability at least*  $1-\delta$ *, we have*

$$
R_u(\hat{\mathbf{w}}_m) - R_u(\mathbf{w}_u^*) \le \frac{N}{u} \left( \frac{c_1 r_m^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{m} \right) + \frac{N}{m} \left( \frac{c_1 r_u^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{u^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{u} \right),
$$

**1492 1493 1494**

**1489 1490 1491**

**1495 1496** *where*  $c_1$ ,  $c_2$  *and*  $c_3$  *are some positive constants.* 

> <span id="page-27-3"></span> $=\frac{u}{u}$ n

**1497 1498 1499 1500** *Proof of Lemma* [9.](#page-27-1) Note that since  $w_u^*$  is also an empirical risk minimizer computed on the test set., the results of Lemma [8](#page-25-1) also hold for  $\mathbf{w}_u^*$  with every  $m$  in the statement replaced by u. Also note that the following holds almost surely:

(27)

 $\Box$ 

$$
0 \leq R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*)
$$
  
=  $R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*) - \hat{R}_m(\hat{\mathbf{w}}_m) + \hat{R}_m(\mathbf{w}_N^*) + \hat{R}_m(\hat{\mathbf{w}}_m) - \hat{R}_m(\mathbf{w}_N^*)$   
 $\leq R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*) - \hat{R}_m(\hat{\mathbf{w}}_m) + \hat{R}_m(\mathbf{w}_N^*)$ 

 $\left(R_u(\hat{\mathbf{w}}_m)-R_u(\mathbf{w}_N^*)-\hat{R}_m(\hat{\mathbf{w}}_m)+\hat{R}_m(\mathbf{w}_N^*)\right)$ 

1507 and  
\n1508  
\n1509  
\n1510  
\n
$$
0 \le R_N(\hat{\mathbf{w}}_u) - R_N(\mathbf{w}_N^*)
$$
\n
$$
= R_N(\hat{\mathbf{w}}_u) - R_N(\mathbf{w}_N^*) - R_u(\hat{\mathbf{w}}_u) + R_u(\mathbf{w}_N^*) + R_u(\hat{\mathbf{w}}_u) - R_u(\mathbf{w}_N^*)
$$
\n
$$
\le R_N(\hat{\mathbf{w}}_u) - R_N(\mathbf{w}_N^*) - R_u(\hat{\mathbf{w}}_u) + R_u(\mathbf{w}_N^*)
$$
\n
$$
= \frac{m}{n} \left( \hat{R}_m(\hat{\mathbf{w}}_u) - \hat{R}_m(\mathbf{w}_N^*) - R_u(\hat{\mathbf{w}}_u) + R_u(\mathbf{w}_N^*) \right),
$$
\n(28)

**1512 1513** where last equations in both cases use the equation  $N \cdot R_N(\mathbf{w}) = m \cdot \hat{R}_m(\mathbf{w}) + u \cdot R_u(\mathbf{w})$ .

**1514 1515 1516** Now we are going to use [\(26\)](#page-27-2) obtained in the proof of Lemma [8.](#page-25-1) Using [\(27\)](#page-27-3) and, subsequently, employing [\(26\)](#page-27-2) for  $f = \ell(\hat{\mathbf{w}}_m; \cdot) - \ell(\mathbf{w}_N^*; \cdot)$ , where we subtract  $E_m f$  for both sides of (26), for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we obtain:

1517 
$$
0 \le R_u(\hat{\mathbf{w}}_m) - R_u(\mathbf{w}_N^*) - \hat{R}_m(\hat{\mathbf{w}}_m) + \hat{R}_m(\mathbf{w}_N^*)
$$
  
\n1518  
\n1519  
\n1520 
$$
\le \frac{N}{u} \left( \inf_{K'>1} \frac{K'}{K'-1} \hat{R}_m(\hat{\mathbf{w}}_m - \mathbf{w}_N^*) + \frac{2(1+\eta)^2 K' \lambda r_m^*}{B} \left( 1 + \frac{1}{K_2 \sqrt{\log \frac{4}{\delta}}} \right)^2 \right)
$$
  
\n1521  
\n1522  
\n1523  $+ 32(1+\beta) \binom{N}{\mathbf{k}'} R \log \frac{24}{\mathbf{k}'} + 4 \frac{CK_2 \sqrt{2 \log \frac{12}{\delta}}}{\mathbf{k}'} + \frac{4K' \lambda^{i-1} \sqrt{2C^2 K \log N \log \frac{24}{\delta}}}{\mathbf{k}'}.$ 

$$
+32(1+\beta)\left(\frac{N}{m^2}\right)K'\lambda B\log\frac{24}{\delta}+4\frac{CK_2\sqrt{2\log\frac{12}{\delta}}}{m}+\frac{4K'\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{24}{\delta}}}{(K-1)m}\Bigg).
$$

**1525 1526 1527** Similarly, the same argument can be used for  $w_u^*$ , which gives that for any  $\delta \in (0,1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we obtain:

$$
0 \leq \hat{R}_m(\hat{\mathbf{w}}_u) - \hat{R}_m(\mathbf{w}_N^*) - R_u(\hat{\mathbf{w}}_u) + R_u(\mathbf{w}_N^*)
$$
  

$$
\leq \frac{N}{m} \left( \inf_{K' > 1} \frac{K'}{K' - 1} R_u(\hat{\mathbf{w}}_u - \mathbf{w}_N^*) + \frac{2(1 + \eta)^2 K' \lambda r_u^*}{B} \left( 1 + \frac{1}{K_2 \sqrt{\log \frac{4}{\delta}}} \right)^2 \right)
$$

**1532 1533 1534**

**1535**

**1539**

**1541**

**1524**

$$
+\ 32(1+\beta)\left(\frac{N}{u^2}\right)K'\lambda B\log\frac{24}{\delta}+4\frac{CK_2\sqrt{2\log\frac{12}{\delta}}}{u}+\frac{4K'\lambda^{i-1}\sqrt{2C^2K\log N\log\frac{24}{\delta}}}{(K-1)u}\Bigg).
$$

**1536 1537** The union bound gives us that both inequalities hold simultaneously with probability at least  $1 - \delta$ , summing these two inequalities, we obtain

1538 
$$
0 \le R_u(\hat{\mathbf{w}}_m) - R_u(\mathbf{w}_u^*) - \hat{R}_m(\hat{\mathbf{w}}_m) + \hat{R}_m(\mathbf{w}_u^*)
$$
  
\n1539 
$$
\le \frac{N}{u} \left( \frac{c_1 r_m^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{m} \right) + \frac{N}{m} \left( \frac{c_1 r_u^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{u^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{u} \right).
$$
  
\n1542

**1543 1544** Using the fact the  $\hat{w}_m$  and  $w_u^*$  are the empirical risk minimizers on the training and test set, respectively, we finally get:

$$
0 \le R_u(\hat{\mathbf{w}}_m) - R_u(\mathbf{w}_u^*)
$$
  
\n
$$
\le \frac{N}{u} \left( \frac{c_1 r_m^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{m} \right) + \frac{N}{m} \left( \frac{c_1 r_u^*}{B \log \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{u^2} + \frac{c_3 \sqrt{\log N \log \frac{12}{\delta}}}{u} \right),
$$
  
\nwhere  $c_n$  and  $c_n$  are some positive constants.

**1549** where  $c_1, c_2$  and  $c_3$  are some positive constants.

**1550 1551** The proof is completed.

**1552 1553**

**1557 1558 1559**

**1561 1562 1563**

**1554 1555 1556** *Proof of Theorem [3.](#page-4-2)* Notice that  $2B\Re_N\{f \in \mathcal{F}^* : Ef^2 \leq r\} \leq \psi_m(r)$ , according to Assump-tion [1,](#page-4-1) we have  $\log \mathcal{N}(\varepsilon, \mathcal{W}, L_2(\mathbb{P})) \leq \mathcal{O}(\log(1/\varepsilon))$ . Using Dudley's integral bound [\[35\]](#page-10-5) to find  $\psi_m$  and solving  $r \leq \mathcal{O}(B\psi_m(r))$ , it is not hard to verify that

$$
r^* \leq \mathcal{O}\left(\frac{B^2\log m}{m}\right).
$$

**1560** Insert the solution r<sup>\*</sup> into Lemma [9,](#page-27-1) for any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , we have

$$
\varepsilon_u(\hat{\mathbf{w}}_m) = \mathcal{O}\Bigg(\frac{N}{mu}\left(\log m + \log u + \frac{N\log\frac{1}{\delta}}{m} + \frac{N\log\frac{1}{\delta}}{u} + \sqrt{\log N\log\frac{1}{\delta}}\right)\Bigg).
$$

**1564 1565** The proof is complete.

 $\Box$ 

 $\Box$ 

**1566 1567 1568** The detailed proof of Theorem [4](#page-4-3) is completely similar with Theorem [3,](#page-4-2) In consequence, we omit here and give the Lemmas for sub-exponential.

**1569 1570 1571** Lemma 10 (Peeling Lemma for sub-exponential). *Assume that there is a constant* B > 0 *such that for every*  $f \in \mathcal{F}^*$  *we have*  $Ef^2 \leq B \cdot Ef$ . Suppose Assumptions *1* hold and the objective function  $\ell(\cdot;\cdot)$  *is sub-exponential. Assume there is a sub-root function*  $\psi_m(r)$  *such that* 

$$
2B\Re_N\{f \in \mathcal{F}^* : Ef^2 \le r\} \le \psi_m(r),
$$

**1573 1574** *where*  $E_{r,m}$  was defined in [\(17\)](#page-22-0). Let  $r_m^*$  be a fixed point of  $\psi_m(r)$ .

**1575 1576** *Fix some*  $\lambda > 1$ *. For*  $w(r, f) = \min\{r\lambda^k : k \in \mathbb{N}, r\lambda^k \ge Ef^2\}$ *, define the following rescaled version of excess loss class:*

**1577**

**1572**

**1578 1579**

**1590 1591 1592**  $\mathcal{G}_r = \left\{\frac{r}{m(r)}\right\}$  $\frac{r}{w(r,f)}f : f \in \mathcal{F}^*$ .

**1580** *Then for any*  $r > r_m^*$  *and*  $t > 0$ *, with probability at least*  $1 - \delta$ *, we have* 

$$
\sup_{g \in \mathcal{G}_r} Eg - E_m g \le \frac{(1+\eta)\sqrt{rr_m^*}}{B} \left(1 + \frac{1}{K_1 \log \frac{2}{\delta}}\right)
$$

$$
+4\sqrt{(1+\beta)\left(\frac{N}{m^{2}}\right)r\log\frac{12}{\delta}}+8\sqrt{\frac{3C^{2}K\log^{2}N}{m^{2}}\log\frac{12}{\delta}}
$$

,

 $\boldsymbol{u}$ 

**1587** *where*  $K, K_1, \eta, \beta$  *are some positive constants. C is positive constants depending on*  $\eta, \beta$ *.* 

**1588 1589 Lemma 11.** *Under the assumptions of Theorem [4,](#page-4-3) for any*  $\delta \in (0,1)$ *, with probability at least*  $1-\delta$ *, we have*

$$
R_N(\hat{\mathbf{w}}_m) - R_N(\mathbf{w}_N^*) \le \frac{c_1 r_m^*}{B \log^2 \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log^2 N \log \frac{12}{\delta}}}{m},
$$

**1593 1594** *where*  $c_1$ ,  $c_2$  *and*  $c_3$  *are some positive constants.* 

**1595 1596 Lemma 12.** *Under the assumptions of Theorem [4,](#page-4-3) for any*  $\delta \in (0,1)$ *, with probability at least*  $1-\delta$ *, we have*

$$
R_u(\hat{\mathbf{w}}_m) - R_u(\mathbf{w}_u^*) \le \frac{N}{u} \left( \frac{c_1 r_m^*}{B \log^2 \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{m^2} + \frac{c_3 \sqrt{\log^2 N \log \frac{12}{\delta}}}{m} \right) + \frac{N}{m} \left( \frac{c_1 r_u^*}{B \log^2 \frac{2}{\delta}} + \frac{c_2 N \log \frac{12}{\delta}}{u^2} + \frac{c_3 \sqrt{\log^2 N \log \frac{12}{\delta}}}{u} \right),
$$

**1601 1602**

**1603**

**1605 1606**

**1604** *where*  $c_1$ ,  $c_2$  *and*  $c_3$  *are some positive constants.* 

### <span id="page-29-1"></span><span id="page-29-0"></span>C.2 SOME RESULTS FOR SUB-EXPONENTIAL FUNCTIONS IN SUBSECTION [4.2](#page-5-0)

 $\overline{m}$ 

**1607 1608 1609 1610 Theorem 11.** *Suppose Assumptions* [2,](#page-6-1) [3,](#page-6-2) [4,](#page-7-0) and [5](#page-7-1) hold. For any  $w \in W$ , let the loss function  $\ell(\mathbf{w};\cdot)$  *be sub-exponential. Suppose that the step sizes*  $\{\eta_t\}$  *satisfies*  $\eta_t = \frac{1}{t+t_0}$  *such that*  $t_0 \ge$  $\max\{(2P)^{1/\alpha}, 1\}$ *. For any*  $\delta \in (0, 1)$ *, with probability*  $1 - \delta$ *,* 

*(a). If*  $\alpha \in (0, \frac{1}{2})$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)T^{\frac{1-2\alpha}{2}}\log\bigg(\frac{1}{\delta}\bigg) + \frac{N}{u}\sqrt{\frac{\log^3\big(\frac{1}{\delta}\big)}{m}}\bigg).
$$

(*b*). If  $\alpha = \frac{1}{2}$ , we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log(T)\log\left(\frac{1}{\delta}\right) + \frac{N}{u}\sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}}\bigg).
$$

**1613 1614 1615**

(c). If 
$$
\alpha \in (\frac{1}{2}, 1]
$$
, we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \mathcal{O}\left(L_{\mathcal{F}}\frac{\sqrt{N}}{u}\log^{\frac{1}{2}}(T)\log\left(\frac{1}{\delta}\right) + \frac{N}{u}\sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}}\right)
$$

.

<span id="page-30-1"></span>.

<span id="page-30-0"></span>**Corollary 2.** *Suppose Assumptions* [2,](#page-6-1) [3,](#page-6-2) [4,](#page-7-0) and [5](#page-7-1) hold. For any  $w \in W$ , let the loss function  $\ell(\mathbf{w};\cdot)$  *be sub-exponential. Suppose that the learning rate*  $\{\eta_t\}$  *satisfies*  $\eta_t = \frac{2}{\mu(t+t_0)}$  *such that*  $t_0 \geq \max\{\frac{2}{\mu}(2P)^{\frac{1}{\alpha}},1\}$ *. For any*  $\delta \in (0,1)$ *, with probability*  $1-\delta$ *,* 

(a). If 
$$
\alpha \in (0, \frac{1}{2})
$$
, we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{Nd}}{u}\log^{\frac{1}{2}}(T)T^{\frac{1}{2}-\alpha}\log\bigg(\frac{1}{\delta}\bigg) + \frac{N}{u}\sqrt{\frac{\log^3\big(\frac{1}{\delta}\big)}{m}} + \frac{1}{T^{\alpha}}\bigg),
$$

(*b*). If  $\alpha = \frac{1}{2}$ , we have

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{Nd}}{u}\log(T)\log\left(\frac{1}{\delta}\right) + \frac{N}{u}\sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}} + \frac{1}{T^{\alpha}}\bigg).
$$

*(c). If*  $\alpha \in (\frac{1}{2}, 1)$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^*) = \mathcal{O}\bigg(L_{\mathcal{F}}\frac{\sqrt{Nd}}{u}\log^{\frac{1}{2}}(T)\log(1/\delta) + \frac{N}{u}\sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}} + \frac{1}{T^{\alpha}}\bigg).
$$

*(d). If*  $\alpha = 1$ *, we have* 

$$
R_u(\mathbf{w}^{(T+1)}) - R_u(\mathbf{w}^*) = s\mathcal{O}\left(L_{\mathcal{F}}\frac{\sqrt{Nd}}{u}\log^{\frac{1}{2}}(T)\log(1/\delta) + \frac{N}{u}\sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}} + \frac{\log(T)\log^3(1/\delta)}{T}\right)
$$

#### **1650** C.3 PROOFS OF SUBSECTION [4.2](#page-5-0)

**1651 1652 1653 1654 1655** *Proof of Theorem [5.](#page-7-2)* In order to obtain high-probability bounds with our new concentration inequalities, for the term  $\sup_{f_{\mathbf{w}} \in \mathcal{F}_{\mathcal{W}}} \sum_{\mathbf{x} \in \mathbf{X}_m} f_{\mathbf{w}}(\mathbf{x}) = \sup_{\mathbf{w} \in \mathcal{W}} \sum_{\mathbf{x} \in \mathbf{X}_m} (R_N(\mathbf{w}) - \ell(\mathbf{w}; \mathbf{x})) =$  $m \cdot \sup_{\mathbf{w}\in\mathcal{W}}(R_N(\mathbf{w}) - R_m(\mathbf{w}))$ , where we obtain a factor of m in the equation because in Theorem [1](#page-2-2) we considered unnormalized sums.

**1656** To use Theorem [1,](#page-2-2) we need to bound  $\left\|\max_\mathbf{x} \sup_{f_\mathbf{w} \in \mathcal{F}_\mathcal{W}} f_\mathbf{w}(\mathbf{x})\right\|$ 2  $\tilde{\psi}_2$ , we have

$$
\begin{array}{c} 1657 \\ 1658 \end{array}
$$

 $\overline{11}$ 

1658  
\n
$$
\left\|\max_{\mathbf{x}} \sup_{f_{\mathbf{w}} \in \mathcal{F}_{\mathcal{W}}} f_{\mathbf{w}}(\mathbf{x})\right\|_{\psi_2}^2 \le \left\|\max_{\mathbf{x}} \sup_{\mathbf{w} \in \mathcal{W}} \ell(\mathbf{w}; \mathbf{x})\right\|_{\psi_2}^2 \le K^2 \max_{\mathbf{x}} \left\|\sup_{\mathbf{w} \in \mathcal{W}} \ell(\mathbf{w}; \mathbf{x})\right\|_{\psi_2}^2 \log N \le K^2 K_2^2 \log N.
$$

**1660 1661 1662 1663** where K and  $K_2$  are two positive constants. The second inequality holds using Theorem [7](#page-12-3) [\[34\]](#page-10-15) and the last inequality satisfies because  $\ell(\cdot; \mathbf{x})$  is sub-Gaussian, using property of the tail bound for sub-Gaussian distribution.

**1664 1665** Then we turn to bound  $\sigma_W^2$ . For any fixed  $\mathbf{w} \in \mathcal{W}$  and any  $\delta \in (0,1)$ , with at least probability  $1-\frac{\delta}{2}$ , we have

$$
\frac{1}{N}\sum_{\mathbf{x}\in\mathbf{Z}_N}(\ell(\mathbf{w};\mathbf{x})-R_N(\mathbf{w}))^2=\frac{1}{N}\sum_{\mathbf{x}\in\mathbf{Z}_N}\ell(\mathbf{w};\mathbf{x})^2-R_N(\mathbf{w})^2\leq \frac{1}{N}\sum_{\mathbf{x}\in\mathbf{Z}_N}\ell(\mathbf{w};\mathbf{x})^2\leq K\log\frac{2}{\delta},
$$

**1669 1670** where K is a positive constant. the last inequality holds because  $\ell(\cdot; \mathbf{x})$  is sub-Gaussian, then  $\ell(\cdot; \mathbf{x})^2$ is sub-exponential, using property of the tail bound for sub-exponential distribution. Thus for any  $\delta \in (0, 1)$ , with at least probability  $1 - \frac{\delta}{2}$ , we have

**1671**

$$
\sigma_W^2 = \sup_{\mathbf{w} \in \mathcal{W}} \left( \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} (\ell(\mathbf{w}; \mathbf{x}) - R_N(\mathbf{w}))^2 \right) \le K \log \frac{2}{\delta}
$$
(29)

**1674 1675 1676** According to Theorem [1,](#page-2-2) Let  $Q_m = m \cdot (R_N(w) - \hat{R}_m(w))$ , and combined with [\(29\)](#page-30-1). For any  $\delta \in (0, 1)$  with probability at least  $1 - \delta$ , we have

$$
\begin{array}{c} 1677 \\ 1678 \end{array}
$$

 $\sup_{\mathbf{w}\in\mathcal{W}}(R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w}))$ 

$$
\frac{1680}{1681}
$$

**1679**

**1683 1684**

**1685 1686**

<span id="page-31-0"></span>
$$
\leq (1+\eta)E_m + 2\sqrt{\left(\frac{4(1+\beta)K\log\frac{2}{\delta}}{m} + \frac{2C^2K^2K_2^2\log n}{m^2}\right)\log\frac{12}{\delta}}
$$
  

$$
\leq (1+\eta)E_m + 4\sqrt{\frac{(1+\beta)K\log\frac{2}{\delta}\log\frac{12}{\delta}}{m} + \frac{2\sqrt{2C^2K^2K_2^2\log N\log\frac{12}{\delta}}}{m}}
$$
(30)  

$$
\leq (1+\eta)E_m + 4\sqrt{\frac{(1+\beta)K}{\log\frac{12}{\delta}} + \frac{2\sqrt{2C^2K^2K_2^2\log N\log\frac{12}{\delta}}}{m}}.
$$

1<sub>2</sub>

$$
\leq (1+\eta)E_m + 4\sqrt{\frac{(1+\beta)K}{m}}\log\frac{12}{\delta} + \frac{2\sqrt{2C^2K^2K_2^2}\log N\log\frac{12}{\delta}}{m}
$$

**1687 1688 1689**

**1690 1691 1692 1693** where the second inequality holds using  $\sqrt{a+b} \leq \sqrt{a}$  + √ b. Next, we need to bound the  $E_m = \mathbb{E} \left[ \sup_{\mathbf{w} \in \mathcal{W}} \left( R_N(\mathbf{w}) - \frac{1}{m} \sum_{i=1}^m \ell(\mathbf{w}; \xi_i) \right) \right]$ . We have

<span id="page-31-1"></span>
$$
E_m = \mathbb{E}\left[\sup_{\mathbf{w}\in\mathcal{W}}\left(R_N(\mathbf{w}) - \frac{1}{m}\sum_{i=1}^m \ell(\mathbf{w}; \xi_i)\right)\right]
$$
  
\n
$$
\leq 2\mathbb{E}_{\xi\sim\mathbf{X}_N, v}\left[\sup_{\mathbf{w}\in\mathcal{W}} v_i\left(R_N(\mathbf{w}) - \frac{1}{m}\sum_{i=1}^m \ell(\mathbf{w}; \xi_i)\right)\right]
$$
  
\n
$$
\leq 2\mathbb{E}_v\left[\sup_{\mathbf{w}\in\mathcal{W}}\sum_{i=1}^m v_i R_N(\mathbf{w})\right] + 2\mathbb{E}_{\xi\sim\mathbf{X}_N, v}\left[\sup_{\mathbf{w}\in\mathcal{W}}\frac{1}{m}\sum_{i=1}^m v_i \ell(\mathbf{w}; \xi_i)\right]
$$
(31)

m

 $i=1$ 

**1701 1702**

**1703**

**1704 1705**

**1706** where the first inequality holds using symmetrization inequality (see Lemma 11.4 [\[3\]](#page-9-7)).

 $i=1$ 

 $= 2\Re R_N(\mathbf{w}),$ 

**1707 1708** Recall that for any  $\hat{w}$ , we have

$$
R_u(\hat{\mathbf{w}}) - \hat{R}_m(\hat{\mathbf{w}}) \leq \frac{N}{u} \sup_{\mathbf{w} \in \mathcal{W}} R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w}).
$$

**1711 1712 1713**

**1716**

**1709 1710**

**1714 1715** Thus, Combining [\(30\)](#page-31-0), [\(31\)](#page-31-1) and above inequality, for any  $\delta \in (0,1)$  with probability at least  $1-\delta$ , we have

<span id="page-31-2"></span>1717  
\n1718  
\n1719  
\n
$$
R_u(\hat{\mathbf{w}}) - \hat{R}_m(\hat{\mathbf{w}}) \le \frac{2N(1+\eta)\Re R_N(\mathbf{w})}{u} + 4\frac{N}{u}\sqrt{\frac{(1+\beta)K}{m}}\log\frac{12}{\delta} + \frac{2N\sqrt{2C^2K^2K_2^2\log N\log\frac{12}{\delta}}}{mu}.
$$
\n(32)

**1720 1721 1722 1723 1724 1725 1726 1727** Next, we need to bound the Rademacher complexity with traditional Dudley's integral technique. Firstly, we denote some notations. Let  $d_{\mathcal{W}}(\mathbf{w}, \mathbf{w}') = \left(\frac{1}{N} \sum_{i=1}^{N} [\ell(\mathbf{w}; \mathbf{x}_i) - \ell(\mathbf{w}'; \mathbf{x}_i)]^2\right)^{\frac{1}{2}}$ . For  $j \in \mathbb{N}$ , let  $\alpha_j = 2^{-j}M$  with  $M = \sup_{\mathbf{w} \in \mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}, \mathbf{w}^{(1)})$ , where  $\mathcal{W}_R$  denotes the parameter space consisting of the initial parameters  $w^{(1)}$  together with all possible  $w^{(i)}$  that can be obtained using Algorithm [1.](#page-6-0) Denote by  $T_j$  the minimal  $\alpha_j$ -cover of  $W_R$  and  $\ell(\mathbf{w}^j; \mathbf{x})[\mathbf{w}]$  the element in  $T_j$ that covers  $\ell(\mathbf{w}; \mathbf{x})$ . Specifically, since  $\{\ell(\mathbf{w}^{(1)}; \mathbf{x})\}$  is a M-cover of  $\mathcal{W}_R$ , we set  $\ell(\mathbf{w}^0; \mathbf{x})[\mathbf{w}] =$  $\ell(\mathbf{w}^{(1)};\mathbf{x})[\mathbf{w}]$ , (Note that  $\mathbf{w}^{(1)}$  is the initialization parameter and  $\mathbf{w}^{j}$  is the associated parameter of **1728 1729**  $\ell$  in  $T_i$ ). For arbitrary  $n \in \mathbb{N}$ :

**1730 1731**

1731  
1732  

$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\sum_{i=1}^N v_i\ell(\mathbf{w};\mathbf{x}_i)\right]
$$

$$
= \mathbb{E}_{\boldsymbol{v}} \left[ \sup_{\mathbf{w} \in \mathcal{W}_R} \bigg( \sum_{i=1}^N \Big( v_i(\ell(\mathbf{w}; \mathbf{x}_i) - \ell(\mathbf{w}^n; \mathbf{x}_i))[\mathbf{w}] \right. \\ \left. n \right]
$$

1

<span id="page-32-3"></span>
$$
+\sum_{j=1}^{n} v_i(\ell(\mathbf{w}^j; \mathbf{x}_i)[\mathbf{w}] - \ell(\mathbf{w}^{j-1}; \mathbf{x}_i)[\mathbf{w}]) + v_i\ell(\mathbf{w}^{(1)}; \mathbf{x}_i)\bigg)\Bigg]
$$
(33)

$$
\leq \mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\left(\sum_{i=1}^N v_i(\ell(\mathbf{w};\mathbf{x}_i)-\ell(\mathbf{w}^n;\mathbf{x}_i)[\mathbf{w}])\right)\right]+\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^N v_i\ell(\mathbf{w}^{(1)};\mathbf{x}_i)\right] + \sum_{i=1}^n \mathbb{E}_{\mathbf{w}}\left[\sup_{\mathbf{x}\in\mathcal{W}}\left(\sum_{i=1}^N v_i(\ell(\mathbf{w}^j;\mathbf{x}_i)[\mathbf{w}]-\ell(\mathbf{w}^{j-1};\mathbf{x}_i)[\mathbf{w}])\right)\right],
$$

$$
+\sum_{j=1}^n\mathbb{E}_{\boldsymbol{v}}\left[\sup_{\boldsymbol{\mathrm{w}}\in\mathcal{W}_R}\bigg(\sum_{i=1}^Nv_i(\ell(\boldsymbol{\mathrm{w}}^j;\boldsymbol{\mathrm{x}}_i)[\boldsymbol{\mathrm{w}}]-\ell(\boldsymbol{\mathrm{w}}^{j-1};\boldsymbol{\mathrm{x}}_i)[\boldsymbol{\mathrm{w}}]\bigg)\bigg)\right].
$$

For the first term, we apply Cauchy-Schwarz inequality and obtain

## **1748 1749**

$$
\begin{array}{c} 1750 \\ 1751 \end{array}
$$

$$
\frac{1752}{175}
$$

**1754**

**1753**

$$
f_{\rm{max}}
$$

<span id="page-32-2"></span>
$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_{R}}\left(\sum_{i=1}^{N}v_{i}(\ell(\mathbf{w};\mathbf{x}_{i})-\ell(\mathbf{w}^{n};\mathbf{x}_{i})[\mathbf{w}])\right)\right]
$$
\n
$$
\leq \left(\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^{N}v_{i}^{2}\right]\right)^{\frac{1}{2}}\left(\sup_{\mathbf{w}\in\mathcal{W}_{R}}\sum_{i=1}^{N}(\ell(\mathbf{w};\mathbf{x}_{i})-\ell(\mathbf{w}^{n};\mathbf{x}_{i})[\mathbf{w}])^{2}\right)^{\frac{1}{2}}\leq N\alpha_{n}.
$$
\n(34)

# By Massart's Lemma, we have

**1759 1760 1761**

<span id="page-32-1"></span>
$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\left(\sum_{i=1}^N v_i(\ell(\mathbf{w}^j;\mathbf{x}_i)[\mathbf{w}]-\ell(\mathbf{w}^{j-1};\mathbf{x}_i)[\mathbf{w}])\right)\right]
$$
\n
$$
\leq \sqrt{N}\sup_{\mathbf{w}\in\mathcal{W}_R}d_{\mathcal{W}}(\mathbf{w}^j,\mathbf{w}^{j-1})\sqrt{2\log|T_j||T_{j-1}|}.
$$
\n(35)

By the Minkowski inequality,

$$
\sup_{\mathbf{w}\in\mathcal{W}_R}d_{\mathcal{W}}(\mathbf{w}^j,\mathbf{w}^{j-1})
$$

$$
\begin{split}\n&= \sup_{\mathbf{w} \in \mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}^j; \mathbf{x}_i) [\mathbf{w}] - \ell(\mathbf{w}; \mathbf{x}) + \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}^{j-1}; \mathbf{x}_i) [\mathbf{w}] \right]^2 \right)^{\frac{1}{2}} \\
&< \sup \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}^j; \mathbf{x}_i) [\mathbf{w}] - \ell(\mathbf{w}; \mathbf{x}) \right]^2 \right)^{\frac{1}{2}}\n\end{split} \tag{36}
$$

<span id="page-32-0"></span>
$$
\leq \sup_{\mathbf{w}\in\mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}^j; \mathbf{x}_i) [\mathbf{w}] - \ell(\mathbf{w}; \mathbf{x}) \right]^2 \right)
$$

1778  
\n1779  
\n1780  
\n
$$
+ \sup_{\mathbf{w}\in\mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}^{j-1}; \mathbf{x}_i) [\mathbf{w}] \right]^2 \right)
$$
\n1781

$$
= \sup_{\mathbf{w}\in\mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}^j, \mathbf{w}) + \sup_{\mathbf{w}\in\mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}, \mathbf{w}^{j-1}) \le \alpha_j + \alpha_{j-1} = 3\alpha_j.
$$

**1782 1783 1784** Plugging [\(36\)](#page-32-0) into [\(35\)](#page-32-1), using facts that  $\alpha_j = 2(\alpha_j - \alpha_{j+1})$  and  $|T_j| \ge |T_{j-1}|$ , taking summation over  $j$ ,

**1785 1786**

**1819 1820**

**1826 1827 1828**

**1830 1831 1832**

**1834 1835**

$$
\sum_{j=1}^{n} \mathbb{E}_{\mathbf{v}} \left[ \sup_{\mathbf{w} \in \mathcal{W}_{R}} \left( \sum_{i=1}^{N} v_{i} (\ell(\mathbf{w}^{j}; \mathbf{x}_{i})[\mathbf{w}] - \ell(\mathbf{w}^{j-1}; \mathbf{x}_{i})[\mathbf{w}]) \right) \right]
$$
\n
$$
\leq 6\sqrt{N} \sum_{j=1}^{n} \alpha_{j} \sqrt{\log |T_{j}|} = 12\sqrt{N} \sum_{j=1}^{n} (\alpha_{j} - \alpha_{j+1}) \sqrt{\log |T_{j}|}
$$
\n
$$
= 12\sqrt{N} \sum_{j=1}^{n} (\alpha_{j} - \alpha_{j+1}) \sqrt{\log \mathcal{N}(\alpha_{j}, \mathcal{W}_{R}, d_{\mathcal{W}})}
$$
\n
$$
\leq 12\sqrt{N} \int^{\alpha_{0}} \sqrt{\log \mathcal{N}(\alpha, \mathcal{W}_{R}, d_{\mathcal{W}})} d\alpha \leq 12\sqrt{N} \int^{\infty} \sqrt{\log \mathcal{N}(\alpha, \mathcal{W}_{R}, d_{\mathcal{W}})} d\alpha.
$$
\n(37)

<span id="page-33-0"></span>
$$
\leq 12\sqrt{N} \int_{\alpha_{n+1}} \sqrt{\log N(\alpha, \mathcal{W}_R, d\mathcal{W})} d\alpha \leq 12\sqrt{N} \int_{\alpha_{n+1}} \sqrt{\log N(\alpha, \mathcal{W}_R, d\mathcal{W})} d\alpha
$$

For the last term, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$  we have

<span id="page-33-1"></span>
$$
\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^{N} v_i \ell(\mathbf{w}^{(1)}; \mathbf{x}_i)\right] \leq \left(\sum_{i=1}^{N} \ell^2(\mathbf{w}^{(1)}; \mathbf{x}_i)\right)^{\frac{1}{2}} \leq K \sqrt{N \log \frac{2}{\delta}},\tag{38}
$$

**1802 1803 1804** where  $K$  is a positive constant. The first inequality holds by Khintchine-Kahane inequality [\[22\]](#page-10-18). The second inequality satisfies because  $\ell(\cdot; \mathbf{x})$  is sub-Gaussian, therefore,  $\ell(\cdot; \mathbf{x})$  is sub-exponential. Using Lemma [4,](#page-12-4) we can derive the inequality.

**1805 1806** Taking the limit as  $n \to \infty$ , plugging [\(34\)](#page-32-2), [\(37\)](#page-33-0) and [\(38\)](#page-33-1) into [\(33\)](#page-32-3) and combining with the difination of Rademacher complexity, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we have

<span id="page-33-2"></span>
$$
\Re R_N(\mathbf{w}) = \frac{1}{N} \mathbb{E}_v \left[ \sup_{\mathbf{w} \in \mathcal{W}} \sum_{i=1}^N v_i \ell(\mathbf{w}; \mathbf{x}_i) \right] \le \frac{K \sqrt{\log \frac{2}{\delta}}}{\sqrt{N}} + \frac{12}{\sqrt{N}} \int_0^\infty \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d\mathcal{W})} d\varepsilon, \tag{39}
$$

**1811 1812** where  $v_i$  is Rademacher random variable. One can verify that  $d_{W_R}(\ell(\mathbf{w};\cdot),\ell(\mathbf{w}';\cdot)) =$  $\max_{z \in \mathcal{Z}} |\ell(\mathbf{w}; z) - \ell(\mathbf{w}'; z)|$  is a metric in  $\mathcal{W}_R$ . we have

> $d_{\mathcal{W}} \leq$  $\sqrt{1}$ N  $\sum_{i=1}^{N}$  $i=1$  $\left[\max_{\mathbf{w},\mathbf{w}'\in\mathcal{W}_R,\mathbf{x}\in\mathcal{Z}}\ell(\mathbf{w};z_i)-\ell(\mathbf{w}';\mathbf{x}_i)\right]^2\right)^{\frac{1}{2}}$  $\leq d_{\mathcal{W}_R}.$

**1817 1818** By the definition of covering number, we have  $\mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}}) \leq \mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}_R})$ . Besides, applying Lemma [1](#page-6-3) yields

$$
d_{\mathcal{W}_{R}} = \max_{\mathbf{x} \in \mathcal{Z}} |\ell(\mathbf{w}; z) - \ell(\mathbf{w}'; z)| \leq L_{\mathcal{F}} \|\mathbf{w} - \mathbf{w}'\|_{2}.
$$

**1821 1822 1823 1824** By the definition of covering number, we have  $\mathcal{N}(\varepsilon,\mathcal{W}_R,d_{\mathcal{W}_R})\leq \mathcal{N}\left(\frac{\varepsilon}{L_\mathcal{F}},\mathcal{B}(\mathbf{w}^{(1)},R),d_{\mathbf{w}}\right)$ , where  $d_{\mathbf{w}}(\mathbf{w}, \mathbf{w}') = \|\mathbf{w} - \mathbf{w}'\|_2$  and  $\mathcal{W}_R \in \mathcal{B}(\mathbf{w}^{(1)}, R)$ .

**1825** According to [\[33\]](#page-10-19),  $\log N\left(\epsilon, \mathcal{B}(\mathbf{w}^{(1)}, R), d_{\mathbf{w}}\right) \le d \log(3R/\epsilon)$  holds. Therefore, we obtain

<span id="page-33-3"></span>
$$
\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}}) \le d \log \left( \frac{3L_{\mathcal{F}}R}{\varepsilon} \right). \tag{40}
$$

**1829** Furthermore,

<span id="page-33-4"></span>
$$
d_{\mathcal{W}}^2(\mathbf{w}, \mathbf{w}^{(1)}) = \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}; \mathbf{x}_i) - \ell(\mathbf{w}^{(1)}; \mathbf{x}_i) \right]^2 \le L_{\mathcal{F}}^2 R^2,
$$

**1833** where the last inequality is due to Lemma [1.](#page-6-3) This implies that

$$
\int_0^\infty \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d\mathcal{W})} \, d\varepsilon = \int_0^{L_{\mathcal{F}}R} \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d\mathcal{W})} \, d\varepsilon. \tag{41}
$$

≤

**1836 1837** Combining [\(39\)](#page-33-2), [\(40\)](#page-33-3), and [\(41\)](#page-33-4), for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$  yields

N

$$
\begin{array}{c} 1838 \\ 1839 \end{array}
$$

<span id="page-34-0"></span>
$$
\mathcal{R}_N(\mathbf{w}) \leq \frac{K\sqrt{\log\frac{2}{\delta}}}{\sqrt{N}} + 12\sqrt{\frac{d}{N}} \int_0^{L_{\mathcal{F}}R} \sqrt{\log(3L_{\mathcal{F}}R/\varepsilon)} d\varepsilon
$$

$$
\leq \frac{K\sqrt{\log\frac{2}{\delta}}}{\sqrt{N}} + 12\sqrt{\frac{d}{N}} \left(\sqrt{\log 3} + \frac{3}{2}\sqrt{\pi}\right) L_{\mathcal{F}}R.
$$

 $\left(\sqrt{\log 3} + \frac{3}{2}\right)$ 

 $\sqrt{\pi}$   $L_{\mathcal{F}}R$ .

(42)

 $\Box$ 

**1840 1841**

**1842**

**1846 1847 1848**

**1853**

**1861 1862**

**1843 1844 1845** Applying Theorem 47 in [\[27\]](#page-10-13) to bound R in [\(42\)](#page-34-0) and plugging in [\(32\)](#page-31-2) with probability  $1 - \delta/2$ , we conclude that with probability at least  $1 - \delta$ ,

 $+12\sqrt{\frac{d}{\lambda}}$ N

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \begin{cases} \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log^{\frac{1}{2}}(T) T^{\frac{1}{2} - \alpha} \log\left(\frac{1}{\delta}\right) + \frac{N \log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right) & \text{If } \alpha \in \left(0, \frac{1}{2}\right) \\ \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log(T) \log\left(\frac{1}{\delta}\right) + \frac{N \log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right) & \text{If } \alpha = \frac{1}{2} \end{cases}
$$

**1849 1850 1851** u δ u 2 O L<sup>F</sup> √ Nd u log 1 <sup>2</sup> (T) log( <sup>1</sup> δ ) + <sup>N</sup> log( 1 δ ) u <sup>√</sup><sup>m</sup> If α ∈ 1 2 , 1 .

**1852** The proof is complete.

**1854 1855 1856 1857** *Proof of Theorem [11.](#page-29-1)* In order to obtain high-probability bounds with out new concentration inequalities, for the term  $\sup_{f_{\mathbf{w}} \in \mathcal{F}_{\mathcal{W}}} \sum_{\mathbf{x} \in \mathbf{X}_m} f_{\mathbf{w}}(\mathbf{x}) = \sup_{\mathbf{w} \in \mathcal{W}} \sum_{\mathbf{x} \in \mathbf{X}_m} (R_N(\mathbf{w}) - \ell(\mathbf{w}; \mathbf{x})) =$  $m \cdot \sup_{w \in \mathcal{W}} (R_N(w) - R_m(w))$ , where we obtain a factor of m in the equation because in Theorem [2](#page-3-0) we considered unnormalized sums.

**1858 1859** Then, to use Theorem [2,](#page-3-0) we need to bound  $\left\|\max_{\mathbf{x}}\sup_{f_{\mathbf{w}}\in\mathcal{F}_{\mathcal{W}}}f_{\mathbf{w}}(\mathbf{x})\right\|$ 2  $\frac{2}{\psi_1}$ .

1860  
\n
$$
\left\|\max_{\mathbf{x}} \sup_{f_{\mathbf{w}} \in \mathcal{F}_{\mathcal{W}}} f_{\mathbf{w}}(\mathbf{x})\right\|_{\psi_1}^2 \le \left\|\max_{\mathbf{x}} \sup_{\mathbf{w} \in \mathcal{W}} \ell(\mathbf{w}; \mathbf{x})\right\|_{\psi_1}^2 \le K^2 \max_{\mathbf{x}} \left\|\sup_{\mathbf{w} \in \mathcal{W}} \ell(\mathbf{w}; \mathbf{x})\right\|_{\psi_1}^2 \log^2 N \le K^2 K_1^2 \log^2 N.
$$

**1863 1864 1865** where K and  $K_1$  are two constants. The second inequality holds using Theorem [7](#page-12-3) [\[34\]](#page-10-15) and the last inequality satisfies because  $\ell(\cdot; \mathbf{x})$  is sub-exponential, using property of the tail bound for subexponential distribution

**1866 1867** Then we turn to bound  $\sigma_W^2$ . For any fixed  $\mathbf{w} \in \mathcal{W}$  and any  $\delta \in (0,1)$ , with at least probability  $1-\frac{\delta}{2}$ , we have

$$
\frac{1868}{1869} \qquad \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} (\ell(\mathbf{w}; \mathbf{x}) - R_N(\mathbf{w}))^2 = \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} \ell(\mathbf{w}; \mathbf{x})^2 - R_N(\mathbf{w})^2 \le \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} \ell(\mathbf{w}; \mathbf{x})^2 \le K \log^2 \frac{2}{\delta},
$$

**1871 1872 1873** where K is a positive constant. the last inequality holds because  $\ell(\cdot; \mathbf{x})$  is sub-exponential. Thus,  $\ell^2(\cdot;{\bf x})$  is sub-Weibull random variable with tail parameter 2, using Lemma [4](#page-12-4) we can derive the last inequality. Thus for any  $\delta \in (0, 1)$ , with at least probability  $1 - \frac{\delta}{2}$ , we have

$$
\sigma_W^2 = \sup_{\mathbf{w} \in \mathcal{W}} \left( \frac{1}{N} \sum_{\mathbf{x} \in \mathbf{Z}_N} (\ell(\mathbf{w}; \mathbf{x}) - R_N(\mathbf{w}))^2 \right) \le K \log^2 \frac{2}{\delta} \tag{43}
$$

**1879 1880**

**1874**

According to Theorem [2,](#page-3-0) Let  $Q_m = m \cdot (R_N(w) - \hat{R}_m(w))$  and combined with [\(43\)](#page-34-1). For any  $\delta \in (0, 1)$  with probability at least  $1 - \delta$ , we have

$$
\sup_{\mathbf{w}\in\mathcal{W}}(R_N(\mathbf{w})-\hat{R}_m(\mathbf{w}))
$$

**1881 1882 1883**

$$
\leq (1+\eta)E_m + 4\sqrt{\left(\frac{(1+\beta)K\log^2\frac{2}{\delta}}{m} + \frac{3C^2K^2K_1^2\log^2 N}{m^2}\right)\log\frac{12}{\delta}}
$$
(44)

**1884 1885 1886**

<span id="page-34-2"></span>
$$
\leq (1+\eta)E_m + 4\sqrt{\frac{(1+\beta)K\log^2\frac{2}{\delta}\log\frac{12}{\delta}}{m}} + \frac{4\sqrt{3C^2K^2K_1^2\log^2 N\log\frac{12}{\delta}}}{m}
$$
 (44)

<span id="page-34-1"></span>.

$$
\begin{array}{c} 1887 \\ 1888 \\ 1889 \end{array}
$$

$$
\leq (1+\eta)E_m + 4\sqrt{\frac{(1+\beta)K\log^3\frac{12}{\delta}}{m}} + \frac{4\sqrt{3C^2K^2K_1^2\log^2 N\log\frac{12}{\delta}}}{m}
$$

**1890 1891 1892 1893 1894 1895 1896 1897 1898 1899 1900 1901 1902** where the second inequality holds using  $\sqrt{a+b} \leq \sqrt{a}$  + b. Next, we need to bound the  $E_m = \mathbb{E} \left[ \sup_{\mathbf{w} \in \mathcal{W}} \left( R_N(\mathbf{w}) - \frac{1}{m} \sum_{i=1}^m \ell(\mathbf{w}; \xi_i) \right) \right]$ . We have  $E_m = \mathbb{E} \Big[$  $\sup_{\mathbf{w}\in\mathcal{W}}\Bigg(R_N(\mathbf{w})-\frac{1}{m}\Bigg)$ m  $\sum_{ }^{\infty}$  $i=1$  $\ell(\mathbf{w}; \xi_i)$  $\setminus$  1  $\leq 2\mathbb{E}_{\xi\sim\boldsymbol{X}_N,v}\left[\sup_{\boldsymbol{\mathrm{w}}\in\mathcal{W}}v_i\right]$  $\sqrt{ }$  $R_N({\bf w})-\frac{1}{\infty}$ m  $\sum_{ }^{\infty}$  $i=1$  $\ell(\mathbf{w}; \xi_i)$  $\setminus$  1  $\leq 2\mathbb{E}_v\left\lceil$ sup w∈W  $\sum_{ }^{\infty}$  $i=1$  $v_iR_N({\bf w})$  $\bigg] + 2 \mathbb{E}_{\xi \sim \boldsymbol{X}_N,v} \left[ \sup_{\boldsymbol{\mathrm{w}} \in \mathcal{W}} \right]$ 1 m  $\sum_{ }^m$  $i=1$  $v_i\ell(\mathbf{w};\xi_i)$ 1  $= 2\Re R_N(\mathbf{w})$ (45)

√

<span id="page-35-0"></span>where the first inequality holds using symmetrization inequality (see Lemma 11.4 [\[3\]](#page-9-7)).

**1905** Recall that for any  $\hat{w}$ , we have

**1903 1904**

**1906 1907 1908**

**1912**

$$
R_u(\hat{\mathbf{w}}) - \hat{R}_m(\hat{\mathbf{w}}) \le \frac{N}{u} \sup_{\mathbf{w} \in \mathcal{W}} R_N(\mathbf{w}) - \hat{R}_m(\mathbf{w}).
$$

**1909 1910** Thus, Combining [\(44\)](#page-34-2), [\(45\)](#page-35-0) and above inequality, for any  $\delta \in (0,1)$  with probability at least  $1-\delta$ , we have

<span id="page-35-3"></span>1911 
$$
R_u(\hat{\mathbf{w}}) - \hat{R}_m(\hat{\mathbf{w}}) \le \frac{2N(1+\eta)\Re R_N}{u} + 4\frac{N}{u}\sqrt{\frac{(1+\beta)K\log^3\frac{12}{\delta}}{m}} + \frac{4N\sqrt{3C^2K^2K_1^2\log^2 N\log\frac{12}{\delta}}}{mu}.
$$
\n1913

**1914 1915 1916 1917 1918 1919 1920 1921 1922** Next, we need to bound the Rademacher complexity with traditional Dudley's integral technique. Let  $d_{\mathcal{W}}(\mathbf{w}, \mathbf{w}') = \left(\frac{1}{N} \sum_{i=1}^{N} [\ell(\mathbf{w}; \mathbf{x}_i) - \ell(\mathbf{w}'; \mathbf{x}_i)]^2\right)^{\frac{1}{2}}$ . For  $j \in \mathbb{N}$ , let  $\alpha_j = 2^{-j}M$  with  $M =$  $\sup_{w \in \mathcal{W}_R} d_{\mathcal{W}}(w, w^{(1)})$ , where  $\mathcal{W}_R$  denotes the parameter space consisting of the initial parameters  $w^{(1)}$  together with all possible  $w^{(i)}$  that can be obtained using Algorithm [1.](#page-6-0) Denote by  $T_j$  the minimal  $\alpha_j$ -cover of  $W_R$  and  $\ell(\mathbf{w}^j; \mathbf{x})[\mathbf{w}]$  the element in  $T_j$  that covers  $\ell(\mathbf{w}; \mathbf{x})$ . Specifically, since  $\{\ell(\mathbf{w}^{(1)};\mathbf{x})\}\)$  is a M-cover of  $\mathcal{W}_R$ , we set  $\ell(\mathbf{w}^0;\mathbf{x})[\mathbf{w}] = \ell(\mathbf{w}^{(1)};\mathbf{x})[\mathbf{w}]$ , (Note that  $\mathbf{w}^{(1)}$  is the initialization parameter and w<sup>j</sup> is the associated parameter of  $\ell$  in  $T_j$ ). For arbitrary  $n \in \mathbb{N}$ :

$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\sum_{i=1}^N v_i\ell(\mathbf{w}; \mathbf{x}_i)\right]
$$

$$
= \mathbb{E}_{\boldsymbol{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\bigg(\sum_{i=1}^N\Big(v_i(\ell(\mathbf{w};\mathbf{x}_i)-\ell(\mathbf{w}^n;\mathbf{x}_i))[\mathbf{w}]\right.\nonumber\\
$$

<span id="page-35-2"></span>
$$
+\sum_{j=1}^{n}v_i(\ell(\mathbf{w}^j;\mathbf{x}_i)[\mathbf{w}]-\ell(\mathbf{w}^{j-1};\mathbf{x}_i)[\mathbf{w}])+v_i\ell(\mathbf{w}^{(1)};\mathbf{x}_i)\bigg)\Bigg)\Bigg]
$$
(47)

 $\overline{1}$ 

$$
\leq \mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\bigg(\sum_{i=1}^N v_i(\ell(\mathbf{w};\mathbf{x}_i)-\ell(\mathbf{w}^n;\mathbf{x}_i)[\mathbf{w}])\bigg)\right]+\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^N v_i\ell(\mathbf{w}^{(1)};\mathbf{x}_i)\right] + \sum_{j=1}^n \mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_R}\bigg(\sum_{i=1}^N v_i(\ell(\mathbf{w}^j;\mathbf{x}_i)[\mathbf{w}]-\ell(\mathbf{w}^{j-1};\mathbf{x}_i)[\mathbf{w}])\bigg)\right].
$$

**1938** For the first term, we apply Cauchy-Schwarz inequality and obtain  $\overline{N}$ 

<span id="page-35-1"></span> $\blacksquare$ 

**1939**

1939  
\n1940  
\n
$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_{R}}\left(\sum_{i=1}^{N}v_{i}(\ell(\mathbf{w};\mathbf{x}_{i})-\ell(\mathbf{w}^{n};\mathbf{x}_{i})[\mathbf{w}])\right)\right]
$$
\n(48)

1943 
$$
\leq \left(\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^N v_i^2\right]\right)^{\frac{1}{2}} \left(\sup_{\mathbf{w}\in\mathcal{W}_R}\sum_{i=1}^N(\ell(\mathbf{w}; \mathbf{x}_i)-\ell(\mathbf{w}^n; \mathbf{x}_i)[\mathbf{w}])^2\right)^{\frac{1}{2}} \leq N\alpha_n.
$$

**1944 1945** By Massart's Lemma, we have

<span id="page-36-1"></span>
$$
\mathbb{E}_{\mathbf{v}}\left[\sup_{\mathbf{w}\in\mathcal{W}_{R}}\left(\sum_{i=1}^{N}v_{i}(\ell(\mathbf{w}^{j};\mathbf{x}_{i})[\mathbf{w}]-\ell(\mathbf{w}^{j-1};\mathbf{x}_{i})[\mathbf{w}])\right)\right]
$$
\n
$$
\leq\sqrt{N}\sup_{\mathbf{w}\in\mathcal{W}_{R}}d_{\mathcal{W}}(\mathbf{w}^{j},\mathbf{w}^{j-1})\sqrt{2\log|T_{j}||T_{j-1}|}.
$$
\n(49)

By the Minkowski inequality,

<span id="page-36-0"></span>
$$
\sup_{\mathbf{w}\in\mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}^j, \mathbf{w}^{j-1})
$$
\n
$$
= \sup_{\mathbf{w}\in\mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}^j; \mathbf{x}_i) [\mathbf{w}] - \ell(\mathbf{w}; \mathbf{x}) + \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}^{j-1}; \mathbf{x}_i) [\mathbf{w}] \right]^2 \right)^{\frac{1}{2}}
$$
\n
$$
\leq \sup_{\mathbf{w}\in\mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}^j; \mathbf{x}_i) [\mathbf{w}] - \ell(\mathbf{w}; \mathbf{x}) \right]^2 \right)^{\frac{1}{2}} + \sup_{\mathbf{w}\in\mathcal{W}_R} \left( \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}; \mathbf{x}) - \ell(\mathbf{w}^{j-1}; \mathbf{x}_i) [\mathbf{w}] \right]^2 \right)^{\frac{1}{2}}
$$
\n
$$
= \sup_{\mathbf{w}\in\mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}^j, \mathbf{w}) + \sup_{\mathbf{w}\in\mathcal{W}_R} d_{\mathcal{W}}(\mathbf{w}, \mathbf{w}^{j-1}) \leq \alpha_j + \alpha_{j-1} = 3\alpha_j.
$$
\n(50)

Plugging [\(50\)](#page-36-0) into [\(49\)](#page-36-1), using facts that  $\alpha_j = 2(\alpha_j - \alpha_{j+1})$  and  $|T_j| \ge |T_{j-1}|$ , taking summation over  $j$ ,

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\n26
$$
\sqrt{N} \sum_{j=1}^{n} \alpha_j \sqrt{\log |T_j|} = 12\sqrt{N} \sum_{j=1}^{n} (\alpha_j - \alpha_{j+1}) \sqrt{\log |T_j|}
$$
  
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\n1974

For the last term, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$  we have

**1980 1981 1982**

**1990 1991 1992**

**1996 1997**

<span id="page-36-2"></span>**1970**

<span id="page-36-3"></span>
$$
\mathbb{E}_{\mathbf{v}}\left[\sum_{i=1}^{N} v_i \ell(\mathbf{w}^{(1)}; \mathbf{x}_i)\right] \leq \left(\sum_{i=1}^{N} \ell^2(\mathbf{w}^{(1)}; \mathbf{x}_i)\right)^{\frac{1}{2}} \leq K\sqrt{N} \log \frac{2}{\delta},\tag{52}
$$

where  $K$  is a positive constant. The first inequality holds by Khintchine-Kahane inequality [\[22\]](#page-10-18). The second inequality satisfies because  $\ell(\cdot; \mathbf{x})$  is sub-exponential, therefore,  $\ell^2(\cdot; \mathbf{x})$  is sub-weibull random variables with parameter 2. Using Lemma [4,](#page-12-4) we can derive the inequality.

**1987 1988 1989** Taking the limit as  $n \to \infty$ , plugging [\(48\)](#page-35-1), [\(51\)](#page-36-2) and [\(52\)](#page-36-3) into [\(47\)](#page-35-2) and combining with the difination of Rademacher complexity, for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$ , we have

<span id="page-36-4"></span>
$$
\Re R_N(\mathbf{w}) = \frac{1}{N} \mathbb{E}_v \left[ \sup_{\mathbf{w} \in \mathcal{W}} \sum_{i=1}^N v_i \ell(\mathbf{w}; \mathbf{x}_i) \right] \le \frac{K \log \frac{2}{\delta}}{\sqrt{N}} + \frac{12}{\sqrt{N}} \int_0^\infty \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d\mathcal{W})} d\varepsilon, \tag{53}
$$

**1993 1994 1995** where  $v_i$  is Rademacher random variable. One can verify that  $d_{W_R}(\ell(\mathbf{w};\cdot),\ell(\mathbf{w}';\cdot)) =$  $\max_{z \in \mathcal{Z}} |\ell(\mathbf{w}; z) - \ell(\mathbf{w}'; z)|$  is a metric in  $\mathcal{W}_R$ . we have

$$
d_{\mathcal{W}} \leq \left(\frac{1}{N}\sum_{i=1}^N\left[\max_{\mathbf{w},\mathbf{w}'\in\mathcal{W}_R,\mathbf{x}\in\mathcal{Z}}\ell(\mathbf{w};z_i)-\ell(\mathbf{w}';\mathbf{x}_i)\right]^2\right)^{\frac{1}{2}} \leq d_{\mathcal{W}_R}.
$$

**1998 1999 2000** By the definition of covering number, we have  $\mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}}) \leq \mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}_R})$ . Besides, applying Lemma [1](#page-6-3) yields

$$
d_{\mathcal{W}_{R}} = \max_{\mathbf{x} \in \mathcal{Z}} |\ell(\mathbf{w}; z) - \ell(\mathbf{w}'; z)| \leq L_{\mathcal{F}} \|\mathbf{w} - \mathbf{w}'\|_{2}.
$$

**2003 2004 2005** By the definition of covering number, we have  $\mathcal{N}(\varepsilon,\mathcal{W}_R,d_{\mathcal{W}_R})\leq\mathcal{N}\left(\frac{\varepsilon}{L_\mathcal{F}},\mathcal{B}(\mathbf{w}^{(1)},R),d_{\mathbf{w}}\right)$ , where  $d_{\mathbf{w}}(\mathbf{w}, \mathbf{w}') = \|\mathbf{w} - \mathbf{w}'\|_2$  and  $\mathcal{W}_R \in \mathcal{B}(\mathbf{w}^{(1)}, R)$ .

**2006** According to [\[33\]](#page-10-19),  $\log N\left(\varepsilon, \mathcal{B}(\mathbf{w}^{(1)}, R), d_{\mathbf{w}}\right) \le d\log(3R/\varepsilon)$  holds. Therefore, we obtain

<span id="page-37-0"></span>
$$
\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d_{\mathcal{W}}) \le d \log \left( \frac{3L_{\mathcal{F}}R}{\varepsilon} \right). \tag{54}
$$

**2010** Furthermore,

**2001 2002**

**2007 2008 2009**

**2011 2012 2013**

**2015 2016 2017**

$$
d_W^2(\mathbf{w}, \mathbf{w}^{(1)}) = \frac{1}{N} \sum_{i=1}^N \left[ \ell(\mathbf{w}; \mathbf{x}_i) - \ell(\mathbf{w}^{(1)}; \mathbf{x}_i) \right]^2 \le L_\mathcal{F}^2 R^2,
$$

**2014** where the last inequality is due to Lemma [1.](#page-6-3) This implies that

$$
\int_0^\infty \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d_W)} \, \mathrm{d}\varepsilon = \int_0^{L_{\mathcal{F}}R} \sqrt{\log \mathcal{N}(\varepsilon, \mathcal{W}_R, d_W)} \, \mathrm{d}\varepsilon. \tag{55}
$$

**2018 2019** Combining [\(53\)](#page-36-4), [\(54\)](#page-37-0), and [\(55\)](#page-37-1), for any  $\delta \in (0, 1)$ , with probability at least  $1 - \frac{\delta}{2}$  yields

<span id="page-37-2"></span><span id="page-37-1"></span>
$$
\mathcal{R}_N(\mathbf{w}) \leq \frac{K \log \frac{2}{\delta}}{\sqrt{N}} + 12\sqrt{\frac{d}{N}} \int_0^{L_{\mathcal{F}}R} \sqrt{\log (3L_{\mathcal{F}}R/\varepsilon)} d\varepsilon
$$
\n
$$
\leq \frac{K \log \frac{2}{\delta}}{\sqrt{N}} + 12\sqrt{\frac{d}{N}} \left(\sqrt{\log 3} + \frac{3}{2}\sqrt{\pi}\right) L_{\mathcal{F}}R.
$$
\n(56)

Applying Theorem 47 in [\[27\]](#page-10-13) to bound R in [\(56\)](#page-37-2) and plugging in [\(46\)](#page-35-3) with probability  $1 - \delta/2$ , we conclude that with probability at least  $1 - \delta$ ,

$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \begin{cases} \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log^{\frac{1}{2}}(T) T^{\frac{1}{2} - \alpha} \log\left(\frac{1}{\delta}\right) + \frac{N}{u} \sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}}\right) & \text{If } \alpha \in \left(0, \frac{1}{2}\right) \\ \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log(T) \log\left(\frac{1}{\delta}\right) + \frac{N}{u} \sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}}\right) & \text{If } \alpha = \frac{1}{2} \\ \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log^{\frac{1}{2}}(T) \log\left(\frac{1}{\delta}\right) + \frac{N}{u} \sqrt{\frac{\log^3\left(\frac{1}{\delta}\right)}{m}}\right) & \text{If } \alpha \in \left(\frac{1}{2}, 1\right]. \end{cases}
$$

The proof is complete.

There is nothing special about the proofs of Corollary [1](#page-8-0) and Corollary [2,](#page-30-0) which simply involve combining Theorem [5](#page-7-2) (or Theorem [11\)](#page-29-1) with an existing optimization result. Here we give the proof of Corollary [1](#page-8-0) as an example.

**2039 2040** *Proof of Corollary [1.](#page-8-0)* By Lemma 43 in [\[27\]](#page-10-13), we have

<span id="page-37-3"></span>
$$
\hat{R}_{m}(\mathbf{w}^{T+1}) - \hat{R}_{m}(\hat{\mathbf{w}}^{*}) = \begin{cases} \mathcal{O}\left(\frac{1}{T^{\alpha}}\right) & \text{if } \alpha \in (0,1) \\ \mathcal{O}\left(\frac{\log(T)\log^{3}(1/\delta)}{T}\right) & \text{if } \alpha = 1. \end{cases}
$$
\n(57)

By Theorem [5,](#page-7-2)

<span id="page-37-4"></span>
$$
R_u(\mathbf{w}^{(T+1)}) - \hat{R}_m(\mathbf{w}^{(T+1)}) = \begin{cases} \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log^{\frac{1}{2}}(T) T^{\frac{1}{2} - \alpha} \log\left(\frac{1}{\delta}\right) + \frac{N \log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right) & \text{If } \alpha \in \left(0, \frac{1}{2}\right) \\ \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log(T) \log\left(\frac{1}{\delta}\right) + \frac{N \log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right) & \text{If } \alpha = \frac{1}{2} \\ \mathcal{O}\left(L_{\mathcal{F}} \frac{\sqrt{Nd}}{u} \log^{\frac{1}{2}}(T) \log\left(\frac{1}{\delta}\right) + \frac{N \log\left(\frac{1}{\delta}\right)}{u\sqrt{m}}\right) & \text{If } \alpha \in \left(\frac{1}{2}, 1\right]. \end{cases}
$$
(58)

**2051** Combing [\(57\)](#page-37-3) and [\(58\)](#page-37-4) yields the result.  $\Box$ 

 $\Box$