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ABSTRACT

011 The merit of Conformal Prediction (CP), as a distribution-free framework for un-
012 certainty quantification, depends on generating prediction sets that are efficient,
013 reflected in small average set sizes, while adaptive, meaning they signal uncer-
014 tainty by varying in size according to input difficulty. A central limitation for
015 deep conformal classifiers is that the nonconformity scores are derived from soft-
016 max outputs, which can be unreliable indicators of how certain the model truly is
017 about a given input, sometimes leading to overconfident misclassifications or un-
018 due hesitation. In this work, we argue that this unreliability can be inherited by the
019 prediction sets generated by CP, limiting their capacity for adaptiveness. We pro-
020 pose a new approach that leverages information from the pre-softmax logit space,
021 using the Helmholtz Free Energy as a measure of model uncertainty and sample
022 difficulty. By reweighting nonconformity scores with a monotonic transformation
023 of the energy score of each sample, we improve their sensitivity to input difficulty.
024 Our experiments with four state-of-the-art score functions on multiple datasets and
025 deep architectures show that this energy-based enhancement improves the adap-
026 tiveness of the prediction sets, leading to a notable increase in both efficiency and
027 adaptiveness compared to baseline nonconformity scores, without introducing any
028 post-hoc complexity.

1 INTRODUCTION

032 Deploying machine learning models in critical, real-world applications requires not just high accu-
033 racy, but also trustworthy uncertainty quantification. Conformal Prediction (CP) has emerged as an
034 effective framework for this challenge (Vovk et al., 2005). It provides a model-agnostic method to
035 construct prediction sets, $C(X)$, that are guaranteed to contain the true class, Y , with a user-specified
036 probability:

$$P(Y \in C(X)) \geq 1 - \alpha.$$

037 This distribution-free guarantee is a significant asset. However, the practical utility of CP depends on
038 the characteristics of these prediction sets. Ideally, they should be **adaptive** and **efficient**: small for
039 inputs that the model finds easy, and appropriately larger for inputs that are difficult or ambiguous.

040 This adaptiveness is governed by the nonconformity score. While many nonconformity scores are
041 designed to produce adaptive sets, they are typically derived from a model’s final softmax probabili-
042 ties. This choice inherits a fundamental weakness, as softmax outputs are often unreliable indicators
043 of a model’s true uncertainty. They can exhibit overconfidence even for misclassified or out-of-
044 distribution (OOD) inputs. Post-hoc calibration helps reduce this issue, but only to a limited extent,
045 as it cannot fully correct the underlying limitations in uncertainty quantification. (Guo et al., 2017;
046 Lee et al., 2018a; Hein et al., 2019). Consequently, the adaptiveness of these scores is by design
047 limited, which can lead to inefficiently large sets for simple inputs, or misleadingly small sets for
048 difficult ones.

049 One approach to improve adaptiveness involves adjusting the score based on an input-specific mea-
050 sure of difficulty, such as the variance of ensemble predictions, the error predicted by an auxil-
051 iary model (Hernández-Hernández et al., 2022), or the variance estimated via Monte-Carlo dropout
052 (MCD) with a neural network (Cortés-Ciriano & Bender, 2019). This principle is related to Nor-
053 malized Conformal Prediction, which has been shown to produce tighter and more informative sets

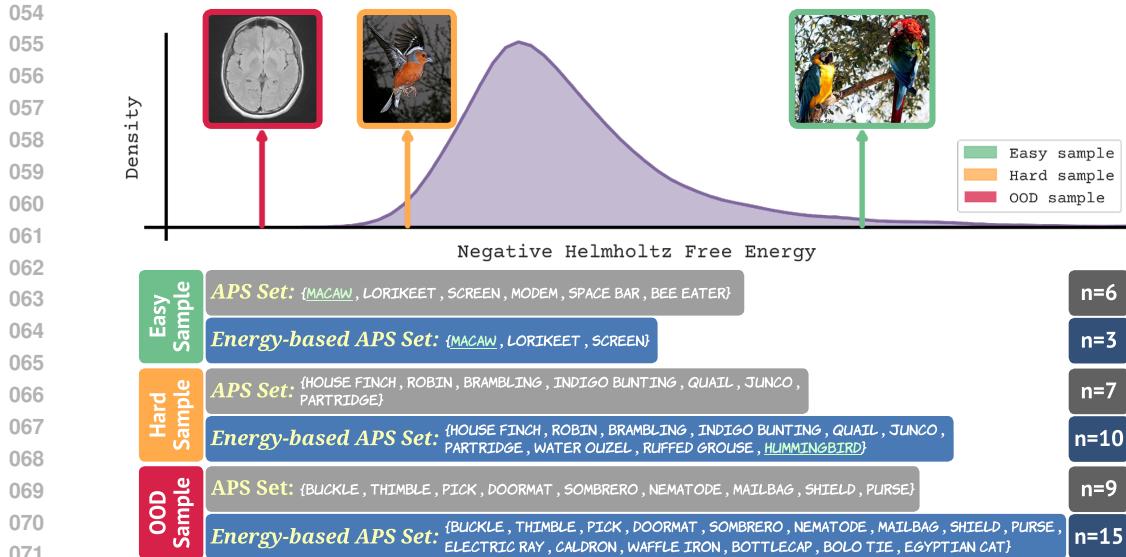


Figure 1: Prediction sets from a standard method (APS [Romano et al. \(2020\)](#)) versus our energy-based variant, demonstrating improved adaptiveness on ImageNet. (i) For an easy input like the image of a *Macaw*, whose clear visual cues (vivid colors, long tail) make it simple to classify, our energy-based method produces a smaller, more efficient set. (ii) For a hard input—a bird image labeled as *Hummingbird*—its appearance deviates from typical hummingbirds (e.g., a thicker, less tapered beak) and shares features with other bird classes, making the image difficult for the model. In this case, the energy-based method returns a larger prediction set, signaling higher uncertainty. (iii) Finally, for an out-of-distribution (OOD) input like a brain MRI that the model was never trained on, our method generates a much larger set, warning the user that the prediction is unreliable. This improvement in adaptive behavior is guided by the Helmholtz free energy, which captures the model’s uncertainty about an input.

in regression by scaling a base score with an uncertainty estimate ([Papadopoulos et al., 2002a; Papadopoulos & Haralambous, 2011; Boström et al., 2017](#)). Building on this principle, we argue that a reliable signal for sample difficulty and model uncertainty exists in the pre-softmax logit space. We propose using the *Helmholtz free energy* computed from the logits, as a principled measure of a model’s familiarity with an input. Inputs aligned with the training data distribution are assigned low energy (high certainty), while atypical or ambiguous inputs receive high energy (low certainty).

The energy signal is then incorporated into the conformal framework by applying per-sample reweighting of the nonconformity scores. For “easy” inputs where the model is certain, the energy-based term magnifies the base score, yielding smaller, more efficient prediction sets. For “hard” or OOD inputs, this term dampens the score, producing larger sets that signal the model’s uncertainty. This improved adaptiveness is exemplified in Figure 1. Compared to existing adaptive scores, our proposed energy-based variants, by leveraging Helmholtz free energy derived from the pre-softmax logit space for per-sample reweighting of nonconformity scores, increase both adaptiveness and efficiency. This approach improves prediction sets across state-of-the-art score functions, while preserving the theoretical coverage guarantees of Conformal Prediction.

We summarize our contributions as follows:

- We provide a theoretical and empirical motivation for moving beyond softmax-based scores. We establish the connection between Helmholtz free energy and model uncertainty and demonstrate that this energy signal distinguishes sample difficulty more effectively than standard softmax metrics.
- We introduce a general framework of **Energy-Based Nonconformity Scores**, which modulates a base nonconformity score with the free energy of each sample to create more adaptive prediction sets that are smaller for easy inputs and larger for difficult or out-of-distribution inputs.
- We provide theoretical and empirical evidence showing that our energy-based enhancement improves the efficiency and adaptiveness of prediction sets for multiple deep learning architectures across a range of different scenarios.

108 **2 MOTIVATION AND METHOD**
 109

110 In this section, we (i) explain why the softmax probabilities that underpin conventional noncon-
 111 formity scores for deep classifiers are often unreliable for efficiently capturing model uncertainty,
 112 (ii) introduce Helmholtz free energy as a more robust measure of uncertainty derived directly from
 113 model logits, and (iii) use this concept to motivate and define a new class of energy-aware scores that
 114 produce more adaptive and efficient prediction sets. All notations are summarized in Appendix A.
 115

116 **2.1 SOFTMAX UNRELIABILITY AND IMPLICATIONS FOR CONFORMAL PREDICTION**
 117

118 Given logits $\mathbf{f}(x)$, a calibrated softmax with temperature $T > 0$ is

$$\hat{\pi}(y | x) = \text{softmax}_y\left(\frac{\mathbf{f}(x)}{T}\right) = \frac{\exp[f_y(x)/T]}{\sum_{k=1}^K \exp[f_k(x)/T]}. \quad (1)$$

121 The common nonconformity scores for classification are functions of $\hat{\pi}$ (as detailed in Ap-
 122 pendix E.2). However, relying on softmax values alone is unreliable for uncertainty assessment.
 123 First, modern networks produce poorly calibrated and often overconfident posteriors (Guo et al.,
 124 2017), including spuriously high confidence on unrecognizable inputs (Nguyen et al., 2015). While
 125 temperature scaling can improve in-distribution calibration, it does not address epistemic uncer-
 126 tainty: OOD, “far” or even “hard” inputs may still map to representation regions that yield confident
 127 softmax outputs (Hein et al., 2019; Lee et al., 2018a). This sensitivity to representation geometry
 128 means that when class manifolds overlap or decision boundaries are poorly separated, softmax con-
 129 fidence can be misleading even after calibration (Cohen et al., 2020). Second, softmax posteriors
 130 entangle likelihoods with learned class priors, biasing scores under label shift or class imbalance.
 131 Margins for minority classes tend to be smaller, intensify uncertainty mis-estimation unless logits
 132 are explicitly adjusted (Ren et al., 2020). Collectively, these issues undermine CP adaptiveness:
 133 probability-based scores can produce (i) unnecessarily large sets for easy samples when tails are
 134 inflated, or (ii) deceptively small sets on ambiguous/OOD inputs that happen to receive high softmax
 135 confidence. For a comprehensive compilation of softmax criticism, see Appendix C.

136 These observations motivate adjusting nonconformity scores with an additional signal that reflects
 137 the model’s holistic signal about its familiarity with x . In the next subsection, we use the *Helmholtz*
 138 *free energy* computed from the logits as a principled, model-aware measure of epistemic uncertainty,
 139 that also correlates with sample difficulty, assigning low energy to easy in-distribution inputs and
 140 high energy to hard, ambiguous, or OOD inputs.

141 **2.2 FREE ENERGY AS A MEASURE OF EPISTEMIC UNCERTAINTY**
 142

143 To quantify a model’s uncertainty in its predictions, we seek a measure that reflects its familiarity
 144 with the input data. We turn to the framework of Energy-Based Models (EBMs) (LeCun et al.,
 145 2006). An EBM defines a scalar *energy* for every configuration of variables, where lower energy
 146 corresponds to higher probability. Any standard discriminative classifier can be interpreted through
 147 the lens of an EBM (Grathwohl et al., 2020). We refer to Appendix D for more details on EBMs.

148 For a classifier with a logit function $f(x) : \mathbb{R}^D \rightarrow \mathbb{R}^K$, we can define a joint energy function over
 149 inputs x and labels y as:

$$E(x, y; f) = -f_y(x), \quad y \in \{1, \dots, K\}. \quad (2)$$

150 This formulation connects the classifier’s outputs directly to an energy landscape. The conditional
 151 probability $p(y|x)$ is then given by the Gibbs-Boltzmann distribution:

$$p(y|x) = \frac{\exp(-E(x, y))}{\sum_{k=1}^K \exp(-E(x, k))} = \frac{\exp(f_y(x))}{\sum_{k=1}^K \exp(f_k(x))}, \quad (3)$$

152 which is identical to the standard softmax function.

153 By marginalizing over the labels, we can derive an unnormalized density over the input space. This
 154 process yields the *Helmholtz free energy*, $F(x)$, which acts as the energy function for the marginal
 155 distribution $p(x)$:

$$F(x; f) = -\tau \log \sum_{k=1}^K \exp\left(\frac{-E(x, k)}{\tau}\right) = -\tau \log \sum_{k=1}^K \exp\left(\frac{f_k(x)}{\tau}\right), \quad (4)$$

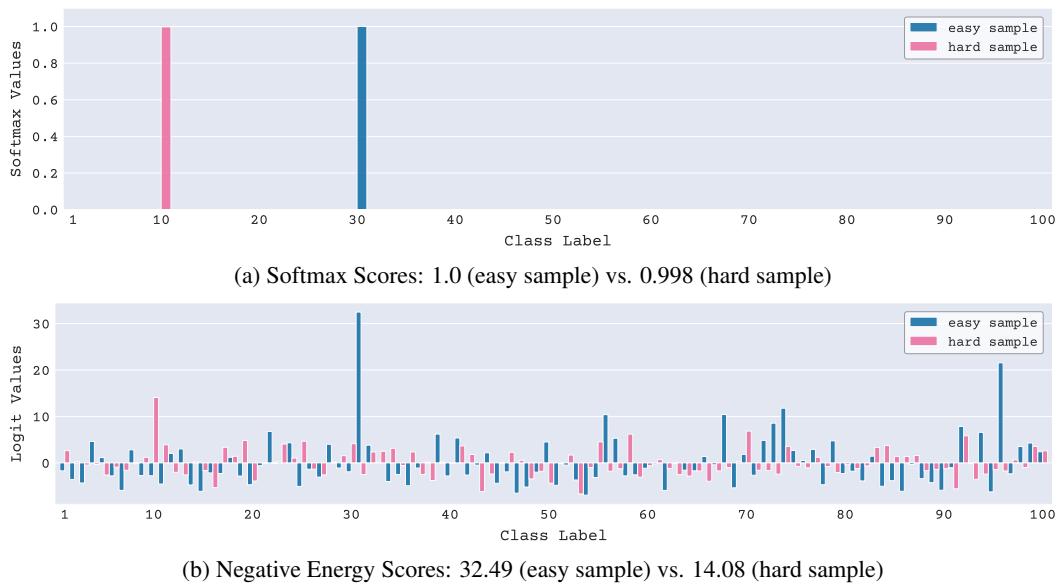


Figure 2: (a) Softmax probability distributions and (b) raw logit outputs of two CIFAR-100 samples computed by a trained ResNet-56. Both samples receive similarly high softmax confidence scores, despite differing significantly in difficulty (1 vs. 27). In contrast, their negative energy scores more clearly reflect this difference.

where τ is a temperature parameter. The free energy represents a soft minimum of the joint energies for a given input x . A low free energy indicates that the model assigns high certainty to at least one class, suggesting the input is familiar. Conversely, a high free energy indicates that the model is uncertain across all classes.

This relationship allows us to define a model-implied marginal density over the input space \mathcal{X} :

$$p(x) = \frac{\exp(-F(x)/\tau)}{Z}, \quad \text{where } Z = \int_{x' \in \mathcal{X}} \exp(-F(x')/\tau) dx', \quad (5)$$

is the partition function, a constant that ensures the distribution integrates to one. This formulation implies that inputs corresponding to high-density regions of the data distribution (i.e., “typical” examples) are assigned low energy, while those in low-density regions have high energy (Liu et al., 2020). We now formalize the connection between free energy and epistemic uncertainty.

Proposition 2.1. *The Helmholtz free energy $F(x)$ is a valid measure of epistemic uncertainty, as it is linearly proportional to the negative log-likelihood of the model-implied data density $p(x)$.*

We refer to Appendix G.2 for proof. This alignment makes the energy score a desirable quantity for epistemic uncertainty (Fuchsgruber et al., 2024; Zong & Huang, 2025) and thus suitable for OOD detection (Liu et al., 2020; Wang et al., 2021).

2.3 ENERGY-BASED NONCONFORMITY SCORES

To clarify our motivation for energy-based conformal classification, we illustrate with a real example how integrating free energy into conformal classification can be beneficial, as it provides additional information not necessarily captured in the softmax space.

Following the definition in Angelopoulos et al. (2021), we quantify the *difficulty* of a sample (x, y_{true}) as

$$D(x, y_{\text{true}}) = o_x(y_{\text{true}}), \quad (6)$$

where $o_x(y_{\text{true}})$ denotes the rank of the true label y_{true} in the model’s predicted class-probability ordering (from most to least likely). Formally,

$$o_x(y) = |\{k \in [K] : \hat{\pi}(k \mid x) \geq \hat{\pi}(y \mid x)\}|. \quad (7)$$

Inspired by the analysis in Liu et al. (2020), Figure 2(a) displays the softmax probability distributions produced by a pretrained ResNet-56 model for two samples from the CIFAR-100 dataset, while Figure 2(b) shows the corresponding raw logit outputs for each sample.

The first sample is considered “easy”, with a difficulty of 1 (i.e., the true label has the highest predicted probability), while the second is “hard”, with a difficulty of 27 (i.e., misclassified by the model). Notably, despite the substantial difference in difficulty, both samples exhibit nearly identical softmax confidence scores, which would make them indistinguishable under standard softmax-based uncertainty metrics. In contrast, their negative energy scores ($-F(x)$), computed from the logits, are significantly more separable. This suggests that $F(x)$ captures a different, and potentially more nuanced, aspect of uncertainty. Easy or high-density samples yield large $-F(x)$, while hard, ambiguous, low-density or OOD samples yield smaller $-F(x)$.

To further investigate this behaviour, Figure 3 shows the distribution of energy scores across the CIFAR-100 test set calculated with a trained ResNet-56 model, stratified by sample difficulty. As the figure illustrates, energy distributions shift noticeably across difficulty levels, suggesting that logits (and their derived energy scores) retain richer information about a model’s confidence than the softmax outputs alone. This highlights energy as an informative signal for uncertainty that can improve the efficiency of nonconformity scores, particularly in cases where softmax probabilities are overconfident or poorly aligned with true sample difficulty.

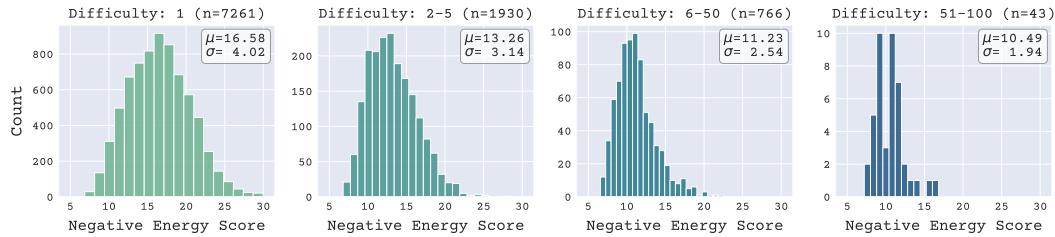


Figure 3: Distribution of negative energy scores ($-F(x)$), stratified by sample difficulty. As difficulty increases, the distribution shifts toward lower energy values, indicating reduced model confidence.

Theorem 2.2 (Monotonicity of Expected Confidence with Sample Difficulty). *Consider two difficulty levels d_1 and d_2 such that $1 \leq d_1 < d_2 \leq K$. Let $\mathbb{E}_{(X,Y) \sim \mathcal{D}}[\cdot | D(X, Y_{true}) = d]$ denote the expectation over the data distribution conditional on samples having difficulty d . For a classifier successfully trained to converge on a representative dataset, the expected negative free energy is a strictly monotonically decreasing function of difficulty:*

$$\mathbb{E}[-F(X) | D(X, Y_{true}) = d_1] > \mathbb{E}[-F(X) | D(X, Y_{true}) = d_2]. \quad (8)$$

The proof of Theorem 2.2 is provided in Appendix G.3.

Having established that the free energy, derived from the logit space, captures epistemic uncertainty and sample difficulty more effectively than softmax probabilities, we propose to integrate the energy score into the nonconformity scores, an approach that aligns with the principles of normalized conformal prediction. By using free energy as a sample-specific difficulty measure, we aim to scale the base nonconformity scores to produce prediction sets that better adjust to the model’s uncertainty regarding each sample.

We define the energy-based variant of a base adaptive nonconformity score, $S(x, y)$, as:

$$S_{\text{Energy-Based}}(x, y) = S(x, y) \cdot \frac{1}{\beta} \log \left(1 + e^{-\beta F(x)} \right). \quad (9)$$

Here, the scaling factor is a softplus function of the negative free energy, $-F(x)$, which ensures a positive, input-dependent weight. The parameter $\beta > 0$ controls the sharpness of this function. This modulation re-calibrates the nonconformity score on a per-sample basis, leveraging the model’s epistemic uncertainty.

The intuition behind this formulation is as follows:

- For “easy” in-distribution samples, the model is certain, resulting in a large negative free energy (i.e., large and positive $-F(x)$). This yields a large scaling factor, which

270 magnifies the base score $S(x, y)$. This reweighting causes the scores of incorrect labels
 271 to more readily exceed the fixed conformal quantile \hat{q} , leading to smaller and more efficient
 272 prediction sets.

273 • **For “hard” or OOD samples**, the model is uncertain, and $-F(x)$ is small or negative. The
 274 scaling factor becomes small, thereby dampening the base score. This dampening reduces
 275 the magnitude of all scores for the given input, causing more plausible labels to fall below
 276 the conformal quantile \hat{q} and thus producing adaptively larger sets that reflect the model’s
 277 uncertainty.

278 As shown in Proposition 2.3, scaling the score is equivalent to adjusting the quantile threshold \hat{q}
 279 for each input, tightening it for confident predictions and relaxing it for uncertain ones. In our
 280 experiments, we apply this modulation to several state-of-the-art scores, including APS, RAPS, and
 281 SAPS. An extension of this modulation to the LAC score is provided in Appendix J.

282 **Proposition 2.3** (Equivalence to Sample-Dependent Thresholding). *Let $S(x, y)$ be any adaptive
 283 base nonconformity score and let $G(x)$ be a positive, sample-dependent scaling function (e.g.,
 284 $G(x) = \text{softplus}(-F(x); \beta) = \frac{1}{\beta} \log(1 + e^{-\beta F(x)})$), assuming this is positive. Let $\mathcal{C}_G(x)$ be the
 285 prediction set constructed using the scaled score $S_G(x, y) = G(x)S(x, y)$ and its corresponding
 286 quantile $\hat{q}_{1-\alpha}^{(G)}$ derived from the calibration set $\mathcal{D}_{\text{cal}} = \{(x_i, y_i)\}_{i=1}^N$.*

287 *This construction is mathematically equivalent to using the original base score $S(x, y)$ with a
 288 sample-dependent threshold $\theta(x)$ that varies for each test sample:*

$$289 \mathcal{C}_G(x) = \{y \in \{1, \dots, K\} \mid S(x, y) \leq \theta(x)\}, \quad (10)$$

290 where the threshold is defined as:

$$291 \theta(x) = \frac{\hat{q}_{1-\alpha}^{(G)}}{G(x)}. \quad (11)$$

292 We refer to Appendix G.4 for proof.

298 3 EXPERIMENTS

300 We present a comprehensive empirical evaluation of our proposed energy-scaled nonconformity
 301 scores, comparing them across various data regimes and distributional challenges. The objective of
 302 CP is to produce prediction sets $\mathcal{C}(X)$ for a test instance X such that its unknown true label Y is
 303 included with a user-specified probability $1 - \alpha$, i.e., $\mathbb{P}(Y \in \mathcal{C}(X)) \geq 1 - \alpha$.

305 3.1 BALANCED TRAINING DATA

306 We first evaluate performance when models are trained on datasets where the prior distribution over
 307 class labels is uniform, i.e., $\mathbb{P}_{\text{train}}(Y = y) = 1/|\mathcal{Y}|$ for all $y \in \mathcal{Y}$. This includes standard ImageNet-
 308 Val, Places365, and CIFAR-100 training sets. For $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$, we report empirical
 309 coverage and average prediction set size in Table 1. This establishes whether energy-based methods
 310 maintain coverage while potentially improving adaptiveness and efficiency under standard, balanced
 311 training conditions. Detailed difficulty-stratified results are also reported in Appendix I.

313 3.2 IMBALANCED TRAINING DATA

315 We then study performance on data with an imbalanced class prior. For this, we use CIFAR-100-LT
 316 training variants, which are designed to simulate long-tailed distributions where class frequencies
 317 decay exponentially ($\mathbb{P}_{\text{train}}(Y = j) \propto \exp(-\lambda \cdot j)$). The parameter λ controls the severity of this
 318 imbalance, with higher values indicating a stronger imbalance, as illustrated in Figure 7.

319 Modern deep networks trained on such long-tailed data exhibit a “familiarity bias”, where the model
 320 shows higher confidence for majority classes and lower confidence for minority classes (Wallace &
 321 Dahabreh, 2012; Samuel et al., 2021). This makes conformal prediction with standard softmax
 322 scores to under-cover minority classes. To address this, our energy-based variants dampen the non-
 323 conformity scores of minority classes more than those of majority classes. This helps to expand
 prediction sets for minority classes, fostering their labels’ inclusion.

324
 325 Table 1: Performance comparison of APS, RAPS, and SAPS nonconformity score functions and their energy-
 326 based variants on CIFAR-100, ImageNet, and Places365 at miscoverage levels $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$.
 327 Results are averaged over 10 trials. For the **Set Size** column, lower is better. **Bold** values indicate the best
 328 performance within each method family (e.g., APS with and without Energy).

Method	$\alpha = 0.1$		$\alpha = 0.05$		$\alpha = 0.025$		$\alpha = 0.01$	
	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size
CIFAR-100 (ResNet-56)								
APS	w/o Energy	0.90 ± 0.01	3.17 ± 0.09	0.95 ± 0.00	6.91 ± 0.24	0.975 ± 0.002	13.29 ± 0.44	0.99 ± 0.00
	w/ Energy	0.90 ± 0.01	3.16 ± 0.08	0.95 ± 0.00	6.49 ± 0.24	0.974 ± 0.001	11.48 ± 0.25	0.99 ± 0.00
RAPS	w/o Energy	0.90 ± 0.00	3.13 ± 0.07	0.95 ± 0.01	8.17 ± 0.47	0.974 ± 0.002	16.38 ± 0.81	0.99 ± 0.00
	w/ Energy	0.90 ± 0.01	3.13 ± 0.08	0.95 ± 0.00	6.18 ± 0.25	0.974 ± 0.002	11.34 ± 0.32	0.99 ± 0.00
SAPS	w/o Energy	0.90 ± 0.01	2.87 ± 0.09	0.95 ± 0.00	7.47 ± 0.43	0.974 ± 0.002	15.08 ± 0.72	0.99 ± 0.00
	w/ Energy	0.90 ± 0.01	2.87 ± 0.11	0.95 ± 0.00	5.94 ± 0.16	0.974 ± 0.001	10.73 ± 0.23	0.99 ± 0.00
ImageNet (ResNet-50)								
APS	w/o Energy	0.90 ± 0.00	1.60 ± 0.02	0.95 ± 0.00	3.99 ± 0.18	0.976 ± 0.001	11.72 ± 0.26	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	1.66 ± 0.03	0.95 ± 0.00	3.84 ± 0.17	0.976 ± 0.001	10.11 ± 0.30	0.99 ± 0.00
RAPS	w/o Energy	0.90 ± 0.00	1.77 ± 0.03	0.95 ± 0.00	4.22 ± 0.06	0.976 ± 0.001	10.56 ± 0.21	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	1.76 ± 0.04	0.95 ± 0.00	3.88 ± 0.07	0.976 ± 0.001	9.18 ± 0.29	0.99 ± 0.00
SAPS	w/o Energy	0.90 ± 0.00	1.67 ± 0.01	0.95 ± 0.00	3.67 ± 0.08	0.976 ± 0.001	9.75 ± 0.31	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	1.66 ± 0.03	0.95 ± 0.00	3.66 ± 0.06	0.976 ± 0.001	8.50 ± 0.29	0.99 ± 0.00
Places365 (ResNet-50)								
APS	w/o Energy	0.90 ± 0.00	7.56 ± 0.13	0.95 ± 0.00	14.28 ± 0.24	0.975 ± 0.002	24.92 ± 0.79	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	7.11 ± 0.11	0.95 ± 0.00	12.98 ± 0.23	0.975 ± 0.002	22.32 ± 0.68	0.99 ± 0.00
RAPS	w/o Energy	0.90 ± 0.00	7.37 ± 0.16	0.95 ± 0.00	14.37 ± 0.27	0.976 ± 0.001	26.34 ± 0.57	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	6.85 ± 0.11	0.95 ± 0.00	12.67 ± 0.23	0.975 ± 0.002	22.35 ± 0.64	0.99 ± 0.00
SAPS	w/o Energy	0.90 ± 0.00	7.20 ± 0.14	0.95 ± 0.00	14.11 ± 0.30	0.976 ± 0.001	25.76 ± 0.49	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	6.79 ± 0.09	0.95 ± 0.00	12.51 ± 0.18	0.975 ± 0.002	22.19 ± 0.69	0.99 ± 0.00

356
 357 Indeed, energy scores capture this training imbalance (Liu et al., 2024a). Figure 4 visually demon-
 358 strates how the distributions of negative energy scores shift across different class bins under both
 359 balanced and imbalanced training conditions.

360 **Theorem 3.1** (Free Energy as an Indicator of Class Imbalance). *Let f be a classifier trained on a
 361 dataset drawn from a distribution $P_{train}(X, Y)$ with imbalanced class priors. Consider two classes,
 362 a majority class y_{maj} and a minority class y_{min} , such that their training priors satisfy $P_{train}(Y = y_{maj}) > P_{train}(Y = y_{min})$.*

364 *Let the model be evaluated on a balanced test distribution P_{test} . Assume the classes are of compara-
 365 ble intrinsic complexity. Then, the expected negative free energy for test samples from the majority
 366 class will be greater than that for the minority class:*

$$\mathbb{E}_{X \sim P_{test}(X|Y=y_{maj})}[-F(X)] > \mathbb{E}_{X \sim P_{test}(X|Y=y_{min})}[-F(X)]. \quad (12)$$

369 See Appendix G.5 for proof. This setup allows us to evaluate how energy influences adaptiveness
 370 when a model’s representations are shaped by imbalanced training. We report marginal coverage
 371 and average set size with the standard balanced CIFAR-100 calibration set and test set, with results
 372 presented in Table 2. We refer to Appendix K for additional experiments on imbalanced scenario.

3.3 RELIABILITY UNDER DISTRIBUTIONAL SHIFT

374 An important test for any uncertainty quantification method is its response to out-of-distribution
 375 (OOD) data. This scenario is particularly challenging for conformal prediction because the assump-
 376 tion of exchangeability between the calibration and test data is violated. Consequently, the formal

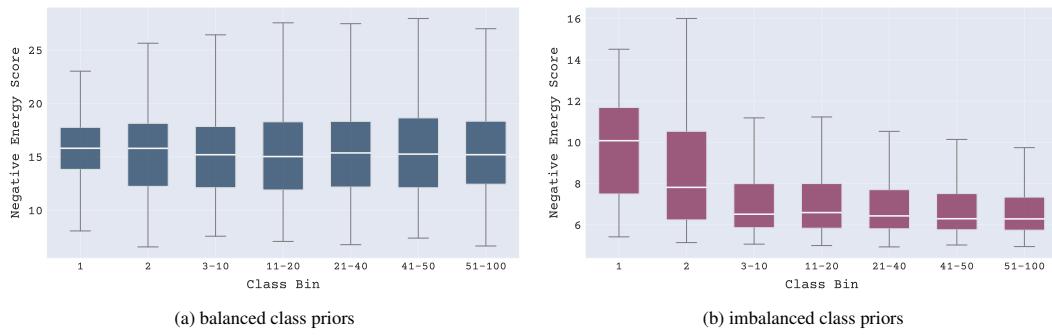


Figure 4: Distributions of negative energy scores across various class bins under balanced and imbalanced training. Results are for CIFAR-100. (a) Balanced model: scores are consistent across class bins. (b) Imbalanced model ($\lambda = 0.03$): minority classes exhibit lower negative energy scores, reflecting reduced confidence.

Table 2: Performance comparison of different nonconformity scores and their energy-based variants on imbalanced CIFAR-100 with an imbalance factor of $\lambda = 0.005$ and at miscoverage levels $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$. Results are averaged over 10 trials with a ResNet-56 model. For the **Set Size** column, lower is better. **Bold** values indicate the best performance within each method family (e.g., AP with and without Energy). Results for additional λ values are provided in Appendix K.

Method	$\alpha = 0.1$		$\alpha = 0.05$		$\alpha = 0.025$		$\alpha = 0.01$		
	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	
CIFAR-100-LT ($\lambda = 0.005$) (ResNet-56)									
APS	w/o Energy	0.90 ± 0.01	8.44 ± 0.30	0.95 ± 0.00	17.22 ± 0.55	0.973 ± 0.003	28.32 ± 0.86	0.99 ± 0.00	45.10 ± 0.75
	w/ Energy	0.90 ± 0.01	7.41 ± 0.22	0.95 ± 0.01	13.30 ± 0.65	0.973 ± 0.004	21.72 ± 1.27	0.99 ± 0.00	34.78 ± 1.72
RAPS	w/o Energy	0.90 ± 0.01	8.88 ± 0.30	0.95 ± 0.00	18.54 ± 0.73	0.972 ± 0.003	30.09 ± 0.94	0.99 ± 0.00	51.65 ± 1.75
	w/ Energy	0.90 ± 0.01	7.59 ± 0.25	0.95 ± 0.01	13.26 ± 0.69	0.973 ± 0.004	22.27 ± 1.24	0.99 ± 0.00	35.43 ± 1.88
SAPS	w/o Energy	0.90 ± 0.01	8.59 ± 0.35	0.95 ± 0.00	17.96 ± 0.66	0.972 ± 0.003	29.25 ± 1.04	0.99 ± 0.00	50.30 ± 1.86
	w/ Energy	0.90 ± 0.01	7.58 ± 0.23	0.95 ± 0.01	13.19 ± 0.68	0.973 ± 0.004	21.99 ± 1.27	0.99 ± 0.00	35.18 ± 1.91

guarantee of marginal coverage no longer holds. This challenge is amplified in real-world deployments where a model, calibrated on in-distribution samples, inevitably encounters novel inputs. These inputs can range from simple *covariate shifts* (e.g., familiar objects in new contexts) to more severe *semantic shifts*, where the inputs belong to classes entirely unseen during training.

In the absence of coverage guarantees, a reliable conformal classifier should not provide a small, incorrect prediction without some indication of its uncertainty. This motivates the following desiderata for the behavior of a conformal predictor $C(\cdot)$ when presented with an OOD input x_{ood} drawn from an OOD distribution P_{ood} , compared to an in-distribution input x_{id} drawn from P_{id} .

Desiderata for a Reliable Conformal Classifier on OOD Data We establish the following desiderata for a conformal predictor’s behavior when encountering OOD data, where the standard exchangeability assumption is violated and coverage guarantees no longer hold.

Desideratum 1 (Adaptive Uncertainty Response). *When faced with an out-of-distribution input, a reliable conformal predictor must adapt its output to signal increased uncertainty. This signal should manifest as either an expansion of the prediction set size or as a principled abstention via an empty set. This response is characterized by one or both of the following outcomes:*

(i) *A significant increase in the probability of abstention:*

$$P_{X \sim P_{\text{ood}}}(C(X) = \emptyset) \gg P_{X \sim P_{\text{id}}}(C(X) = \emptyset) \approx 0$$

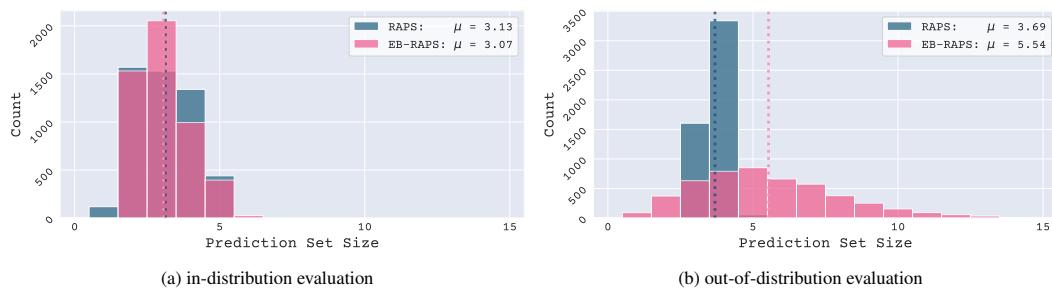
(i) *An inflation in the size of non-empty prediction sets, such that the expected size of non-empty OOD sets is greater than the expected size of in-distribution sets:*

$$\mathbb{E}_{X \sim P_{\text{ood}}}[|C(X)|] > \mathbb{E}_{X \sim P_{\text{id}}}[|C(X)|]$$

432
 433 **Desideratum 2** (Avoidance of False Confidence). *The predictor should minimize the probability of*
 434 *producing a small, non-empty set (e.g., of size 1 or 2) for an OOD input.*

$$435 \quad \text{minimize } P_{X \sim P_{\text{OOD}}}(1 \leq |C(X)| \leq k) \text{ for small } k$$

436 In summary, as also discussed in Appendix P, a larger or empty set is an informative and appropriate
 437 outcome in this scenario, whereas a small, incorrect set is problematic. To assess how our method
 438 aligns with the OOD desiderata, we designed an experiment under a semantic shift. A ResNet-56
 439 model was calibrated on in-distribution CIFAR-100 data and evaluated on the Places365 as the OOD
 440 dataset. As coverage is not a meaningful metric in this context, our analysis focuses on prediction
 441 set size.



452 Figure 5: Prediction set size distributions for the SAPS score and its energy-based variant with $\alpha = 0.05$, on
 453 (a) in-distribution CIFAR-100 and (b) out-of-distribution Places365 data. The energy-based variant produces
 454 larger prediction sets on OOD data. Here, μ represents the overall set size.

455 The results demonstrate an alignment with our desiderata. As shown in Table 3, energy-based
 456 scores produce larger average sets compared to their base counterparts. Figure 5 provides a visual
 457 illustration of this adaptive behavior, comparing the RAPS score with its energy-based variant. The
 458 Energy-based RAPS produces smaller prediction sets on ID data and larger prediction sets for OOD
 459 samples. This response, shows improvement towards Desideratum 1, compared to baseline RAPS.

460 Table 3: Comparison of average prediction set sizes for a ResNet-56 model trained on CIFAR-100. The model
 461 is evaluated on both in-distribution (CIFAR-100) and out-of-distribution (Places365) data. Energy-based vari-
 462 ants demonstrate adaptiveness to the distributional shift by maintaining small sets on ID data while producing
 463 significantly larger sets for OOD inputs. **Bold** values indicate the preferred result: the smallest average set size
 464 for ID (efficiency) and the largest for OOD (uncertainty awareness).

Method	$\alpha = 0.1$		$\alpha = 0.05$		
	Set Size ID (in distribution)	Set Size OOD (out of distribution)	Set Size ID (in distribution)	Set Size OOD (out of distribution)	
RAPS	w/o Energy	3.17 ± 0.09	6.18 ± 0.25	6.91 ± 0.24	14.91 ± 0.81
	w/ Energy	3.16 ± 0.08	86.76 ± 0.94	6.49 ± 0.24	93.40 ± 0.53
SAPS	w/o Energy	3.13 ± 0.07	3.70 ± 0.04	8.17 ± 0.47	8.95 ± 0.47
	w/ Energy	3.13 ± 0.08	5.53 ± 0.07	6.18 ± 0.25	9.05 ± 0.49

4 CONCLUSION

480 This paper demonstrates that the reliability of conformal classifiers is enhanced by moving beyond
 481 softmax probabilities to leverage information about model uncertainty from the logit space. Our
 482 proposed energy-based framework adjusts standard nonconformity scores on a per-sample basis,
 483 leveraging this principled measure of model certainty to make each score sensitive to the model’s
 484 confidence in that specific input. Our evaluations on common nonconformity scores, across multiple
 485 datasets and architectures, confirm that our approach yields prediction sets with improved efficiency
 and adaptiveness, all while preserving the theoretical coverage guarantee.

486

5 REPRODUCIBILITY STATEMENT

488 The empirical results presented in this paper are fully reproducible. Our implementation, based on
489 PyTorch and leveraging the TorchCP library, will be made publicly available. Detailed descriptions
490 of hyperparameters, and environment specification for running experiments, are provided in §F.
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1134 APPENDIX
11351136 A NOTATION
11371138 Table 4: Notation used in this work.
1139

1140 1141	Symbol	Meaning
1142	$x \in \mathbb{R}^D$	Input (feature) vector
1143	K	Total number of classes
1144	$y \in \{1, \dots, K\}$	Class label
1145	$\mathbf{f}(x) = (f_1, \dots, f_K)$	Pre-softmax logit vector produced by the classifier
1146	$\hat{\pi}(y x)$	Model’s softmax probability for class y , Equation 1
1147	$f_{\max}(x)$	$\max_k f_k(x)$
1148	$S(x, y)$	General nonconformity score
1149	T	Temperature used in the calibrated softmax
1150	τ	Temperature used in the energy calculation
1151	α	Desired miscoverage level / target error rate
1152	$\hat{q}_{1-\alpha}$	Quantile threshold for prediction set construction
1153	$E(x, y)$	Joint energy, $E(x, y) = -f_y(x)$
1154	$F(x)$	Helmholtz free energy score, Equation 4
1155	β	Softplus sharpness parameter
1156	$D(x, Y_{\text{true}})$	Sample difficulty measure
1157	$o_x(y)$	Rank of label y in the model’s predicted class-probability ordering
1158	$\mathbb{H}(x)$	Shannon entropy of $\hat{\pi}(y x)$
1159	$\mathcal{D} = \{(x_i, y_i)\}_{i=1}^N$	Dataset with N samples, inputs x_i and labels y_i
1160	$\delta(x)$	Geometric distance of x from the decision boundary

1168 B RELATED WORKS
1169

1170 CP is a statistical framework that provides distribution-free, finite-sample coverage guarantees
1171 for predictions (Vovk et al., 2005). This robust approach to uncertainty quantification has seen
1172 widespread adoption across numerous real-world applications. These include regression (Lei &
1173 Wasserman, 2014; Romano et al., 2019), classification (Sadinle et al., 2019), structured prediction
1174 (Bates et al., 2021), large language models (LLMs) (Su et al., 2024; Cherian et al., 2024; Ku-
1175 mar et al., 2023; Ren et al., 2023; Quach et al., 2024), and diffusion models (Horwitz & Hoshen,
1176 2022; Teneggi et al., 2023), graph neural networks (GNNs) (Zargarbashi et al., 2023; Huang et al.,
1177 2023; Wijegunawardana et al., 2020; Clarkson, 2023; Song et al., 2024), and image generative mod-
1178 els (Horwitz & Hoshen, 2022). Further applications are found in robotic control (Kang et al., 2024;
1179 Luo & Zhou, 2024), hyperspectral imaging (Liu et al., 2024b), healthcare (Lindemann et al., 2024),
1180 finance (Bellotti, 2021), autonomous systems and automated vehicles (Lindemann et al., 2024;
1181 Zecchin et al., 2024; Bang et al., 2024), human-in-the-loop decision making (Straitouri et al., 2023;
1182 Cresswell et al., 2024), bioprocessing (Pham et al., 2025), and scientific machine learning (Moya
1183 et al., 2024; Podina et al., 2024).

1184 The foundational inductive conformal prediction framework (Vovk et al., 2005) often employs a
1185 split conformal (or inductive conformal) approach, where the dataset is divided into a training set
1186 for model fitting and a disjoint calibration set for uncertainty quantification (Papadopoulos et al.,
1187 2002b; Vovk et al., 2005; Shafer & Vovk, 2008; Angelopoulos & Bates, 2021; Lei et al., 2015). This
1188 is done to manage the computational aspects of CP. Beyond this common split, other CP variants

1188 include methods based on cross-validation (Vovk, 2015) or the jackknife (i.e., leave-one-out) technique
 1189 (Barber et al., 2021). The primary goals within CP research are to enhance the efficiency of
 1190 prediction sets (i.e., reduce their size) and to ensure and improve the validity of coverage rates.
 1191

1192 **Improving Prediction Set Efficiency.** Strategies to improve the efficiency of prediction sets pre-
 1193 dominantly fall into two categories: training-time modifications and post-hoc adjustments.
 1194

1195 One line of research focuses on developing new training algorithms or regularizations to learn
 1196 models that inherently produce smaller prediction sets while maintaining coverage (Bellotti, 2021;
 1197 Colombo & Vovk, 2020; Chen et al., 2021; Stutz et al., 2022; Einbinder et al., 2022; Bai et al., 2022;
 1198 Fisch et al., 2021; Yang & Kuchibhotla, 2021; Correia et al., 2024). For example, the uncertainty-
 1199 aware conformal loss aims to optimize APS (Romano et al., 2020) by encouraging non-conformity
 1200 scores towards a uniform distribution (Einbinder et al., 2022), while ConfTr introduces a regular-
 1201 ization term to minimize average set size (Stutz et al., 2022). However, such training methods can be
 1202 computationally intensive due to model retraining and optimization complexity. Early stopping has
 1203 also been explored as a technique to select models leading to more compact prediction sets under
 1204 guaranteed coverage (Liang et al., 2023).

1205 The other avenue involves post-hoc techniques applied to pre-trained models. This includes the de-
 1206 sign of novel non-conformity score functions. Notable examples are LAC (Sadine et al., 2019),
 1207 APS (Romano et al., 2020), RAPS (Angelopoulos et al., 2021), SAPS (Huang et al., 2024b),
 1208 Top-K (Angelopoulos et al., 2021; Luo & Zhou, 2024), and others (Ghosh et al., 2023). Post-
 1209 hoc learning methods have also been proposed (Xi et al., 2024). Research also addresses unique
 1210 settings such as federated learning (Lu et al., 2023; Plassier et al., 2023), multi-label classifica-
 1211 tion (Cauchois et al., 2021; Fisch et al., 2022; Papadopoulos, 2022), outlier detection (Bates et al.,
 1212 2023; Guan & Tibshirani, 2022), and out-of-distribution (OOD) detection (Chen et al., 2023; Nov-
 1213 ello et al., 2024). A common challenge for some post-hoc methods is their reliance on potentially
 1214 unreliable probability outputs from models.

1215 Recent efforts have sought to make conformal prediction sets more adaptive by explicitly incorpo-
 1216 rating epistemic uncertainty. One line of work proposes methods that operate on richer, second-order
 1217 predictions. For instance, Javanmardi et al. (2025) introduce Bernoulli Prediction Sets (BPS), which
 1218 construct provably optimal (i.e., smallest) prediction sets under the assumption that the true data
 1219 distribution is contained within a given "credal set" derived from models like deep ensembles or
 1220 Bayesian neural networks. A complementary, model-agnostic approach is taken by Cabezas et al.
 1221 (2025) with EPICSCORE, which enhances any standard nonconformity score by training a separate
 1222 Bayesian model to learn its conditional distribution. This transforms the score to reflect epistemic
 1223 uncertainty, achieving asymptotic conditional coverage. However, our energy-based framework im-
 1224 proves the adaptiveness of Conformal Classifiers by leveraging uncertainty information from pre-
 1225 softmax logits via the Helmholtz free energy, thus avoiding the need for second-order predictors or
 1226 additional post-hoc computational costs.

1227 **Ensuring Validity and Enhancing Coverage Rates.** A significant body of work is dedicated to
 1228 ensuring the validity of the marginal coverage rate and improving it, particularly under challeng-
 1229 ing conditions, as well as striving for stronger conditional coverage guarantees (Shi et al., 2013;
 1230 Löfström et al., 2015). Efforts have been made to maintain marginal coverage by adapting CP
 1231 to scenarios involving adversarial examples (Gendler et al., 2022; Kang et al., 2024), covariate
 1232 shift (Tibshirani et al., 2019; Deng et al., 2023), label shift (Podkopaev & Ramdas, 2021; Plassier
 1233 et al., 2023), and noisy labels (Feldman et al., 2023; Sesia et al., 2023).

1234 Beyond marginal coverage, many CP algorithms pursue forms of conditional coverage (Vovk, 2012).
 1235 This includes training-conditional validity, which aims to ensure that most training dataset reali-
 1236 zations result in valid marginal coverage on future test data (Bian & Barber, 2023; Pournaderi & Xiang,
 1237 2024). Group-conditional CP methods seek to guarantee coverage across predefined groups within
 1238 the population (Javanmard et al., 2022; Gibbs et al., 2025; Melki et al., 2023). While achieving exact
 1239 pointwise conditional coverage is known to be impossible in general (Foygel Barber et al., 2021),
 1240 practical approaches for approximate or class-conditional coverage exist. For instance, LAC demon-
 1241 strates the possibility of efficient class-conditional coverage (Sadine et al., 2019). Clustered CP
 1242 improves class-conditional coverage by leveraging the label space taxonomy, particularly when the
 1243 number of classes is large (Ding et al., 2023a). Other methods, like k -Class-conditional CP, cali-

bate class-specific score thresholds based on top- k errors. The goal remains to enhance conditional coverage properties while still producing efficient and informative prediction sets.

C ON THE LIMITATIONS OF SOFTMAX OUTPUTS FOR MEASURING MODEL UNCERTAINTY

Our motivation for developing energy-based nonconformity scores, as detailed throughout Section 2, is the inadequacy of softmax outputs for reliably quantifying model uncertainty. To provide a broader context, we reproduce a selection of established criticisms originally compiled in the appendix of Pearce et al. (2021), along with a few more recent perspectives published thereafter.

- “[The softmax output] is often erroneously interpreted as model confidence.” (Gal & Ghahramani, 2016)
- “Deterministic models can capture aleatoric uncertainty but cannot capture epistemic uncertainty.” (Gal et al., 2017)
- “NNs . . . until recently have been unable to provide measures of uncertainty in their predictions.” (Malinin & Gales, 2018)
- “When asked to predict on a data point unlike the training data, the NN should increase its uncertainty. There is no mechanism built into standard NNs to do this . . . standard NNs cannot estimate epistemic uncertainty.” (Pearce, 2020)
- “NNs are poor at quantifying predictive uncertainty.” (Lakshminarayanan et al., 2017)
- “Deep neural networks with the softmax classifier are known to produce highly overconfident posterior distributions even for such abnormal samples.” (Lee et al., 2018b)
- “The output of the [softmax] classifier cannot identify these [far from the training data] inputs as out-of-distribution.” (Hein et al., 2019)
- “The only uncertainty that can reliably be captured by looking at the softmax distribution is aleatoric uncertainty.” (van Amersfoort et al., 2020)
- “Softmax entropy is inherently inappropriate to capture epistemic uncertainty.” (Mukhoti et al., 2021)
- “For [softmax] classifiers . . . misclassification will occur with high confidence if the unknown is far from any known data.” (Boult et al., 2019)
- “. . . softmax output only reflects the total predictive uncertainty instead of the model uncertainty, leading to false confidence under distribution shift.” (Wang et al., 2024)
- “. . . the raw softmax output is neither very reliable . . . nor can it represent all sources of uncertainty” (Gawlikowski et al., 2023)
- “Furthermore, the softmax output cannot be associated with model uncertainty.” (Gawlikowski et al., 2023)
- “. . . the softmax output is often erroneously interpreted as model confidence. In reality, a model can be uncertain in its predictions even with a high softmax output.” (Mobiny et al., 2021)

Additional commentary and remarks also reinforce this view:

- “*Softmax is not telling you anything about . . . model uncertainty.*” — Elise Jennings, Training Program on Extreme-Scale Computing (2019)¹
- “*The [softmax] network has no way of telling you ‘I’m completely uncertain about the outcome and don’t rely on my prediction’.*” — Florian Wilhelm, PyData Berlin (2019)²
- “*Just adding a softmax activation does not magically turn outputs into probabilities.*” — Tucker Kirven, Neural Network Prediction Scores are not Probabilities (2020)³

¹https://youtu.be/Puc_ujh5QZs?t=1323

²<https://youtu.be/LCDIgI-8bHs?t=262>

³https://jtuckerck.github.io/prediction_probabilities.html

1296 **D ENERGY-BASED MODELS**
 1297

1298 An Energy-Based Model (EBM) defines a probability distribution over an input space \mathbb{R}^D through
 1299 an energy function $E_\theta : \mathbb{R}^D \rightarrow \mathbb{R}$, which is typically parameterized by a neural network with
 1300 parameters θ . For any input vector $x \in \mathbb{R}^D$, the probability density is given by the Boltzmann
 1301 distribution:

$$1302 \quad p_\theta(x) = \frac{\exp(-E_\theta(x))}{Z_\theta}, \quad (13)$$

1303 where $Z_\theta = \int_{x'} \exp(-E_\theta(x')) dx'$ is the partition function. This normalization constant is a key
 1304 challenge in EBMs, as its computation involves integrating over the entire high-dimensional input
 1305 space, which is generally intractable.

1306 This framework can be extended to model a joint distribution over inputs and class labels, $p(x, y)$.
 1307 A standard discriminative classifier, which produces a logit vector $\mathbf{f}(x)$, can be re-interpreted as an
 1308 EBM by defining a joint energy function $E(x, y)$. Following common practice [LeCun et al. \(2006\)](#);
 1309 [Grathwohl et al. \(2020\)](#), we define the joint energy as the negative logit corresponding to class y :

$$1311 \quad E(x, y) = -f_y(x). \quad (14)$$

1312 From this joint model, the conditional probability $p(y | x)$ can be derived as:

$$1313 \quad p(y | x) = \frac{p(x, y)}{p(x)} = \frac{p(x, y)}{\sum_{k=1}^K p(x, k)} = \frac{\exp(-E(x, y))}{\sum_{k=1}^K \exp(-E(x, k))} = \frac{\exp(f_y(x))}{\sum_{k=1}^K \exp(f_k(x))}, \quad (15)$$

1314 which is precisely the standard softmax probability $\hat{\pi}(y | x)$. The marginal probability $p(x)$ can
 1315 then be associated with a “free energy” function $E(x) = -\frac{1}{\tau} \log \sum_{k=1}^K \exp(f_k(x)/\tau)$, where τ is
 1316 a temperature parameter.

1317 Training EBMs often proceeds via Maximum Likelihood Estimation (MLE), which aims to shape
 1318 the energy function $E_\theta(x)$ such that it assigns low energy to data points from the true distribution
 1319 and high energy elsewhere. The objective is to maximize the log-likelihood of the observed data \mathcal{D} :

$$1320 \quad \arg \max_{\theta} \mathbb{E}_{x \sim p_{\text{data}}} [\log p_\theta(x)]. \quad (16)$$

1321 The gradient of the log-likelihood with respect to the parameters θ is given by:

$$1322 \quad \nabla_{\theta} \log p_{\theta}(x) = \nabla_{\theta} (-E_{\theta}(x) - \log Z_{\theta}) \quad (17)$$

$$1323 \quad = -\nabla_{\theta} E_{\theta}(x) - \frac{1}{Z_{\theta}} \nabla_{\theta} \int_{x'} \exp(-E_{\theta}(x')) dx' \quad (18)$$

$$1324 \quad = -\nabla_{\theta} E_{\theta}(x) - \int_{x'} \frac{\exp(-E_{\theta}(x'))}{Z_{\theta}} (-\nabla_{\theta} E_{\theta}(x')) dx' \quad (19)$$

$$1325 \quad = -\nabla_{\theta} E_{\theta}(x) + \mathbb{E}_{x' \sim p_{\theta}} [\nabla_{\theta} E_{\theta}(x')]. \quad (20)$$

1326 Remarkably, this gradient can be computed without explicitly evaluating the intractable partition
 1327 function Z_θ . Updating the parameters via stochastic gradient ascent on the log-likelihood is equiva-
 1328 lent to descending on the following loss function:

$$1329 \quad \mathcal{L}_{\text{MLE}} = \mathbb{E}_{x \sim p_{\text{data}}} [E_{\theta}(x)] - \mathbb{E}_{x' \sim p_{\theta}} [E_{\theta}(x')]. \quad (21)$$

1330 This objective can be intuitively understood as a force that “pulls down” the energy of “positive”
 1331 samples drawn from the data distribution (p_{data}) while “pushing up” the energy of “negative” samples
 1332 synthesized from the model’s current distribution (p_θ).

1333 To operationalize this training procedure, we must be able to draw samples x' from the model dis-
 1334 tribution $p_\theta(x)$. Since direct sampling is infeasible, this is typically approximated using Markov
 1335 Chain Monte Carlo (MCMC) methods. A prevalent choice is Stochastic Gradient Langevin Dyna-
 1336 mics (SGLD) [Welling & Teh \(2011\)](#), which iteratively refines an initial sample x_0 (e.g., drawn from a
 1337 simple noise distribution or a buffer of previous samples) according to the rule:

$$1338 \quad x_{t+1} = x_t - \alpha_t \nabla_x E_\theta(x_t) + \sqrt{\eta_t} \epsilon, \quad \text{where } \epsilon \sim \mathcal{N}(0, I), \quad (22)$$

1339 where α_t is the step size and η_t controls the scale of the injected Gaussian noise. After a sufficient
 1340 number of steps, the resulting sample x_T is treated as an approximate sample from $p_\theta(x)$. Different
 1341 strategies for initializing and running the MCMC chain lead to various training algorithms, such as
 1342 Contrastive Divergence (CD) [Hinton \(2002\)](#), which re-initializes the chain from data points at each
 1343 step, and Persistent Contrastive Divergence (PCD) [Tieleman \(2008\)](#), which maintains a persistent
 1344 chain across training iterations to obtain higher-quality samples.

1350 **E EXPERIMENTAL PRELIMINARIES**
 1351

1352 This section details the experimental design, including the conformal prediction framework, the
 1353 nonconformity scores used, and evaluation metrics used to validate our proposed Energy-based non-
 1354 conformity scores.
 1355

1356 **E.1 CONFORMAL PREDICTION FRAMEWORK**
 1357

1358 We employ the standard Split Conformal Prediction (CP) framework in our experiments. A base
 1359 model is trained on a proper training set, and its outputs on a held-out calibration set are used to
 1360 compute non-conformity scores and determine the quantile threshold \hat{q} required to form prediction
 1361 sets. The procedure is formally outlined in Algorithm 1.
 1362

1363 **Algorithm 1** Split Conformal Prediction

1364 1: **Input:** Dataset \mathcal{D} , desired error rate $\alpha \in (0, 1)$, non-conformity score function $S(x, y)$.
 1365 2: Partition \mathcal{D} into a training set $\mathcal{D}_{\text{train}}$ and a calibration set \mathcal{D}_{cal} , such that $\mathcal{D}_{\text{train}} \cap \mathcal{D}_{\text{cal}} = \emptyset$. Let
 1366 $n = |\mathcal{D}_{\text{cal}}|$.
 1367 3: Train the classifier on $\mathcal{D}_{\text{train}}$ to learn the mapping $x \mapsto \mathbf{f}(x)$.
 1368 4: For each example $(x_i, y_i) \in \mathcal{D}_{\text{cal}}$, compute the non-conformity score $s_i = S(x_i, y_i)$.
 1369 5: Calculate the quantile threshold \hat{q} from the set of calibration scores $\{s_1, \dots, s_n\}$. Specifically,
 1370 \hat{q} is the $\frac{\lceil (n+1)(1-\alpha) \rceil}{n}$ -th empirical quantile of these scores.
 1371 6: **Output:** For a new input x_{new} , the prediction set is constructed as:
 1372

1373
$$C(x_{\text{new}}) = \{y \in \{1, \dots, K\} : S(x_{\text{new}}, y) \leq \hat{q}\}$$

1377 **E.2 NONCONFORMITY SCORES FOR DEEP CLASSIFIERS**

1378 The non-conformity score function $S(x, y)$ measures how poorly the label y fits the input x according
 1379 to the trained model. The prediction set $C(x_{\text{new}})$ is then formed by including all labels whose
 1380 non-conformity scores do not exceed the calibrated threshold \hat{q} . A key property of this procedure is
 1381 that the resulting set is guaranteed to contain the true label with a probability of at least $1 - \alpha$, assum-
 1382 ing the test and calibration data points are exchangeable. Exchangeability is a statistical assumption
 1383 that the joint distribution of the data is invariant to permutation, making it a suitable assumption for
 1384 scenarios like simple random sampling.
 1385

1386 We evaluate a range of established non-conformity scores, each based on a different principle for
 1387 measuring how much a model’s prediction disagrees with a given label, as summarized in Table 5.
 1388 We compare these established baselines against their Energy-based counterparts, as defined in Equa-
 1389 tion 9.
 1390

1391 Table 5: Nonconformity scores considered in this work. All scores are functions of $\hat{\pi}(y | x)$ where $u \sim U[0, 1]$;
 1392 The function $o_x(y)$ returns the rank position of label y among all possible labels, ordered by the model’s
 1393 predicted probabilities (with rank 1 being the most likely); $\hat{\pi}_{\max}(x) = \max_k \hat{\pi}(k | x)$; $(\cdot)^+$ denotes the
 1394 positive part; λ and k_{reg} are hyperparameters.

	Method	Nonconformity Score
Adaptive Scores	LAC / THR (Sadinle et al., 2019)	$S_{\text{LAC}}(x, y) = 1 - \hat{\pi}(y x) \equiv -\hat{\pi}(y x)$
	APS (Romano et al., 2020)	$S_{\text{APS}}(x, y) = \sum_{k=1}^K \hat{\pi}(k x) \mathbb{I}\{\hat{\pi}(k x) > \hat{\pi}(y x)\} + u \cdot \hat{\pi}(y x)$
	RAPS (Angelopoulos et al., 2021)	$S_{\text{RAPS}}(x, y) = S_{\text{APS}}(x, y) + \lambda (o_x(y) - k_{\text{reg}})^+$
	SAPS (Huang et al., 2024b)	$S_{\text{SAPS}}(x, y) = \begin{cases} u \cdot \hat{\pi}_{\max}(x), & o_x(y) = 1, \\ \hat{\pi}_{\max}(x) + (o_x(y) - 2 + u) \lambda, & \text{otherwise,} \end{cases}$

1403 Below, we briefly describe how each method works:

1404 **Least Ambiguous Class (LAC/THR)** (Sadinle et al., 2019) is one of the simplest and earliest scores.
 1405 Its non-conformity is defined as $S_{\text{LAC}}(x, y) = 1 - \hat{\pi}(y \mid x)$. The score is thus inversely proportional
 1406 to the model’s confidence; a high probability for the true class yields a low non-conformity score.
 1407 It is worth noting that using the negative probability, $S(x, y) = -\hat{\pi}(y \mid x)$, is mathematically
 1408 equivalent for constructing the prediction set, as the “+1” in the original formula merely shifts the
 1409 range from $[-1, 0]$ to $[0, 1]$ for non-negative interpretability without altering the relative ordering,
 1410 where higher scores indicate greater nonconformity (less conformity between label and sample).
 1411 This is because the conformal procedure relies on the rank-ordering of scores to determine the
 1412 quantile \hat{q} , and the transformation from $1 - \hat{\pi}$ to $-\hat{\pi}$ is monotonic, preserving the rank order. For the
 1413 Energy-based counterpart of this score, we use $-\pi$ (without the bias term). This method provably
 1414 yields the smallest expected prediction sets compared to other methods proposed after this model,
 1415 while preserving the marginal coverage, assuming the predicted probabilities are correct. However,
 1416 this score is non-adaptive, meaning it tends to produce prediction sets of similar size regardless of
 1417 the sample’s intrinsic difficulty, which opened the room for adaptive nonconformity scores and their
 1418 variants to be proposed later.

1419 **Adaptive Prediction Sets (APS)** (Romano et al., 2020) introduced the concept of adaptiveness to
 1420 conformal prediction. The score $S_{\text{APS}}(x, y)$ is the cumulative probability mass of all classes deemed
 1421 more likely than class y . Mathematically, this is the sum of softmax probabilities for all labels k
 1422 whose probability $\hat{\pi}(k \mid x)$ is greater than $\hat{\pi}(y \mid x)$, plus a randomized term to handle ties. This
 1423 design has a crucial effect: for “easy” examples where the model is confident (i.e., $\hat{\pi}_{\max}(x)$ is high
 1424 and Entropy is low), the scores for incorrect labels grow rapidly, leading to small prediction sets.
 1425 Conversely, for “hard” examples where the model is uncertain (a flatter softmax distribution), the
 1426 scores grow slowly, resulting in larger, more inclusive sets that reflect this uncertainty.

1427 **Regularized Adaptive Prediction Sets (RAPS)** (Angelopoulos et al., 2021) builds directly upon
 1428 APS by adding a regularization term. Its score is $S_{\text{RAPS}}(x, y) = S_{\text{APS}}(x, y) + \lambda(o_x(y) - k_{\text{reg}})^+$,
 1429 where $o_x(y)$ is the rank of label y ’s probability. This term penalizes the inclusion of labels with
 1430 a low rank (i.e., large $o_x(y)$), effectively preventing the prediction sets from becoming excessively
 1431 large, especially for uncertain inputs. The hyperparameters k_{reg} and λ control the onset and strength
 1432 of this size-regularizing penalty.

1433 **Sorted Adaptive Prediction Sets (SAPS)** (Huang et al., 2024b) is a more recent refinement that
 1434 aims to mitigate the effects of probability miscalibration in the softmax tail. It treats the top-ranked
 1435 class differently from all others. For labels not ranked first, the score is based on the maximum
 1436 probability $\hat{\pi}_{\max}(x)$ plus a penalty that increases linearly with the label’s rank, weighted by a hy-
 1437 perparameter λ . This approach avoids summing many small, potentially noisy tail probabilities (as
 1438 APS does) and instead relies on the more stable top probability and the rank ordering.

E.3 EVALUATION METRICS

1440 We assess the performance of all methods using a target miscoverage level $\alpha \in$
 1441 $\{0.01, 0.025, 0.05, 0.1\}$. Let $\{(x_i, y_i)\}_{i=1}^{n_{\text{test}}}$ be the test set. The primary metrics are:

- 1443 • **Empirical Coverage:** The fraction of test samples where the true label is included in the
 1444 prediction set.

$$1445 \text{Coverage} = \frac{1}{n_{\text{test}}} \sum_{i=1}^{n_{\text{test}}} \mathbb{I}[y_i \in \mathcal{C}(x_i)] \quad (23)$$

- 1446 • **Macro-Coverage (MacroCov):** While empirical coverage reflects marginal reliability
 1447 over the entire test distribution, MacroCov measures the average per-class coverage, giving
 1448 each class equal weight regardless of its frequency. Let $\hat{c}_y = \frac{1}{|I_y|} \sum_{i \in I_y} \mathbb{I}[y_i \in \mathcal{C}(x_i)]$
 1449 denote the empirical coverage for class y , where $I_y = \{i : y_i = y\}$. Then,

$$1450 \text{MacroCov} = \frac{1}{K} \sum_{y=1}^K \hat{c}_y. \quad (24)$$

1451 This metric is particularly informative in imbalanced or long-tailed settings, since it pre-
 1452 vents head classes from dominating the overall coverage and highlights systematic under-
 1453 coverage of rare classes.

1458 • **Average Prediction Set Size:** The mean size of the prediction sets over the test data.
 1459

1460
$$\text{Size} = \frac{1}{n_{\text{test}}} \sum_{i=1}^{n_{\text{test}}} |\mathcal{C}(x_i)| \quad (25)$$

 1461
 1462

1463 To assess class-conditional reliability, we report the following metrics to verify that they are main-
 1464 tained while reducing the average prediction set size:
 1465

1466 • **Average Class Coverage Gap (CovGap):** This metric measures the average absolute de-
 1467 viation of per-class coverage from the target coverage level $1 - \alpha$ (Ding et al., 2023b). Let
 1468 $I_y = \{i : y_i = y\}$ be the indices of test samples for class y . The empirical coverage for
 1469 class y is $\hat{c}_y = \frac{1}{|I_y|} \sum_{i \in I_y} \mathbb{I}[y_i \in \mathcal{C}(x_i)]$. The gap is then:
 1470

1471
$$\text{CovGap} = \frac{1}{K} \sum_{y=1}^K |\hat{c}_y - (1 - \alpha)| \quad (26)$$

 1472
 1473

1474 We report this value as a percentage.
 1475

1476 To measure the adaptiveness of the prediction sets, we use:
 1477

1478 • **Size-Stratified Coverage Violation (SSCV):** As introduced by Angelopoulos et al. (2021),
 1479 SSCV quantifies whether coverage is maintained across different prediction set sizes. We
 1480 define disjoint set-size strata $\{S_j\}_{j=1}^s$ and group test indices into bins $\mathcal{J}_j = \{i : |\mathcal{C}(x_i)| \in$
 1481 $S_j\}$. The SSCV is the maximum deviation from the target coverage across all bins:
 1482

1483
$$\text{SSCV} = \sup_j \left| \frac{|\{i \in \mathcal{J}_j : y_i \in \mathcal{C}(x_i)\}|}{|\mathcal{J}_j|} - (1 - \alpha) \right| \quad (27)$$

 1484

1485 F REPRODUCIBILITY DETAILS

1487 To ensure full reproducibility, we detail our experimental setup, key hyperparameters, and imple-
 1488 mentation. All source code will be made publicly available. All experiments are implemented based
 1489 on the **TorchCP** library (Huang et al., 2024a), which provides a robust framework for conformal
 1490 prediction on deep learning models. The pre-trained backbone models are sourced from **TorchVi-**
 1491 **sion** (maintainers & contributors, 2016).
 1492

1493 F.1 COMPUTATIONAL ENVIRONMENT

1494 • **Operating System:** Linux kernel 5.14.0-427.42.1.el9_4.x86_64.
 1495 • **GPU Hardware:** NVIDIA H100 80GB HBM3.
 1496 • **CPU Hardware:** 8 cores.
 1497 • **System Memory:** 32 GB RAM.
 1498 • **NVIDIA Driver Version:** 550.144.03.
 1499 • **CUDA Version:** 12.2.
 1500 • **Python Version:** 3.9.21.
 1501 • **PyTorch Version:** 2.0.0+ (with CUDA support).
 1502

1503 F.2 DATASETS AND MODELS

1504 Our experiments are conducted on several standard image classification benchmarks: **CIFAR-100**
 1505 (Krizhevsky et al., 2009), **ImageNet-Val** (Deng et al., 2009), and **Places365** (Zhou et al.,
 1506 2018). These benchmarks are chosen because they contain a large number of classes, which makes
 1507 performance differences between methods more evident. We use pre-trained ResNet (He et al.,
 1508 2016), VGG (Simonyan & Zisserman, 2015), ViT (Dosovitskiy et al., 2021), Swin Transformer (Liu
 1509 et al., 2021), EfficientNet (Tan & Le, 2019), and ShuffleNet (Zhang et al., 2018) architectures from
 1510 1511

1512 TorchVision as our base classifiers. To evaluate performance under distributional shift, we use a
 1513 model trained on CIFAR-100 and test its out-of-distribution (OOD) performance on the Places365
 1514 dataset.

1515 To investigate the methods’ robustness to class imbalance, we create four long-tailed variants
 1516 of CIFAR-100, denoted as **CIFAR-100-LT**. The number of training samples for class $j \in$
 1517 $\{1, \dots, 100\}$, denoted n_j , is set to be proportional to $\exp(-\lambda \cdot j)$. The imbalance factor $\lambda \in$
 1518 $\{0.005, 0.01, 0.02, 0.03\}$ controls the severity of the class imbalance, with larger values of λ cre-
 1519 ating a more pronounced long-tail distribution. For evaluation, we have considered two scenarios:
 1520 (i) the calibration and test sets are balanced while the training data remain imbalanced, and (ii) the
 1521 calibration and test sets follow the same imbalance ratios as the training data.

1522

1523 F.3 HYPERPARAMETER SETTINGS

1524

1525 In our experiments, we set $k_{\text{reg}} = 2$ and $\lambda = 0.2$ for RAPS and we use $\lambda = 0.2$ for SAPS. The soft-
 1526 max probabilities $\hat{\pi}(y|x)$ used by all scores are computed with a temperature parameter T , while
 1527 the free energy $F_{\tau}(x)$ is calculated with its own temperature τ . Crucially, to ensure a fair compar-
 1528 ison, the softmax temperature T was tuned for all baseline and proposed methods to optimize their
 1529 performance. We tune the temperature energy parameter τ with $\ln(\tau) \in [-9, 9]$ and the calibration
 1530 temperature $T \in \{0.01, \dots, 25\}$.

1531 *Remark F.1.* A critical consideration in our proposed modulation is the positivity of the reweight-
 1532 ing factor. When reweighting a base nonconformity score, it is critical the scaling factor must be
 1533 strictly positive. A negative factor would reverse the score’s ordering, invalidating the fundamental
 1534 assumption of conformal prediction that lower scores indicate higher conformity. While our uncer-
 1535 tainty signal is the negative free energy, $-F(x)$, it is not guaranteed to be positive. Mathematically,
 1536 $-F(x)$ becomes negative if $\sum_{k=1}^K \exp(f_k(x)/\tau) < 1$, a condition which implies that the maximum
 1537 logit $f_{\max}(x)$ is negative (a necessary, though not always sufficient, condition). This scenario sig-
 1538 nifies extreme model uncertainty, where the model lacks evidence for any class and typically occurs
 1539 only for far out-of-distribution inputs.

1540

1541 Although we empirically observe that $-F(x)$ is positive for nearly all in-distribution and OOD
 1542 samples in our experiments, to ensure the theoretical robustness of our method, we scale the base
 1543 scores by the softplus of the negative free energy. The hyperparameter β in the softplus function,
 1544 $\frac{1}{\beta} \log(1 + e^{\beta z})$, controls its approximation to the Rectified Linear Unit (ReLU) function. By choos-
 1545 ing a large value for β , the scaling factor softplus($-F(x)$) behaves almost identically to $-F(x)$
 1546 when it is positive, but smoothly transitions to a value near zero in the rare cases where $-F(x) < 0$.
 1547 This behavior is highly beneficial for conformal prediction. For such uncertain inputs, the near-zero
 1548 scaling factor drives the modulated scores for all labels toward zero, causing most or all of them
 1549 to fall below the conformal quantile \hat{q} . This correctly produces an extensively large prediction set,
 1550 signaling the model’s high epistemic uncertainty. Throughout our experiments, we set $\beta = 1$. See
 1551 Appendix O for a detailed ablation study on β .

1552

1553 F.4 CODE AVAILABILITY

1554

1555 The full codebase will be made publicly available to facilitate direct reproduction.

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1566 **G PROOFS**
 1567

1568 **G.1 THEORETICAL VALIDITY OF ENERGY-BASED SCORES**
 1569

1570 In conformal prediction, the validity of the coverage guarantee relies on the exchangeability of the
 1571 nonconformity scores. Specifically, for a set of exchangeable data points, their corresponding non-
 1572 conformity scores must also be exchangeable. This property ensures that the prediction sets contain
 1573 the true label with the desired probability. We now establish that our proposed Energy-modulated
 1574 scores, as defined in Equation 9, satisfy this critical property under standard assumptions. This en-
 1575 sures that using these modulated scores yields valid prediction sets with the guaranteed marginal
 1576 coverage central to conformal prediction theory (Vovk et al., 2005).

1577 **Theorem G.1** (Exchangeability of Energy-Based Nonconformity Scores). *Let $(X_i, Y_i)_{i=1}^{n+1}$ be an
 1578 exchangeable sequence of random variables drawn from a distribution P_{XY} . Assume that:*

1579 (i) *The base nonconformity score $S(x, y)$ is a deterministic function of its arguments.*
 1580
 1581 (ii) *The free energy $F(x)$ is a deterministic function of x as defined in Equation 4.*
 1582

1583 Define the modulated score for $i = 1, \dots, n + 1$ as:

1584
 1585
$$S'_i = S_{\text{Energy-based}}(X_i, Y_i) := S(X_i, Y_i) \cdot \frac{1}{\beta} \log \left(1 + e^{-\beta F(X_i)} \right). \quad (28)$$

 1586

1587 Then the sequence of modulated scores (S'_1, \dots, S'_{n+1}) is exchangeable.
 1588

1589 *Proof.* A sequence of random variables is exchangeable if its joint distribution is invariant under
 1590 any finite permutation of its indices.
 1591

1592 Let the transformation be defined as $h(x, y) = S(x, y) \cdot \frac{1}{\beta} \log \left(1 + e^{-\beta F(x)} \right)$. By assumptions (i)
 1593 and (ii), the base score $S(x, y)$ and the energy function $F(x)$ are deterministic. Since the softplus
 1594 function and multiplication are also deterministic operations, the entire mapping $h(x, y)$ is deter-
 1595 ministic and measurable.

1596 A fundamental property of exchangeable sequences is that they remain exchangeable after applying
 1597 a measurable transformation. That is, if (Z_i) is an exchangeable sequence and g is a measurable
 1598 function, then the sequence $(g(Z_i))$ is also exchangeable.

1599 Applying this principle with $Z_i = (X_i, Y_i)$ and the transformation $g = h$, we find that the sequence
 1600 of scores $(S'_i) = (h(X_i, Y_i))$ inherits exchangeability from the data sequence $((X_i, Y_i))$. Formally,
 1601 for any permutation σ of $\{1, \dots, n + 1\}$,

1603
$$(S'_{\sigma(1)}, \dots, S'_{\sigma(n+1)}) \stackrel{d}{=} (S'_1, \dots, S'_{n+1}), \quad (29)$$

 1604

1605 where $\stackrel{d}{=}$ denotes equality in distribution. Hence, the sequence $(S'_i)_{i=1}^{n+1}$ is exchangeable. \square
 1606

1607 **G.2 PROOF OF PROPOSITION 2.1**
 1608

1609 *Proof.* Let the epistemic uncertainty $U_E(x)$ be defined as the negative logarithm of the model-
 1610 induced input density, $U_E(x) = -\log p(x)$. This definition captures the intuition that uncertainty is
 1611 high where the model assigns low probability density.

1612 Starting from the definition of the input density in Equation 5:

1614
$$p(x) = \frac{\exp(-F(x)/\tau)}{Z}.$$

 1615
 1616

1617 Taking the logarithm of both sides yields:

1618
$$\log p(x) = \log(\exp(-F(x)/\tau)) - \log Z$$

 1619
$$\log p(x) = -F(x)/\tau - \log Z.$$

1620 Multiplying by -1 and rearranging for $F(x)$, we obtain:
 1621

$$\begin{aligned} 1622 \quad -\log p(x) &= F(x)/\tau + \log Z \\ 1623 \quad F(x) &= \tau(-\log p(x)) - \tau \log Z. \end{aligned}$$

1624 Substituting $U_E(x) = -\log p(x)$ and letting $C = -\tau \log Z$ (a constant with respect to x), we arrive
 1625 at:
 1626

$$F(x) = \tau \cdot U_E(x) + C.$$

1627 This shows that the free energy $F(x)$ is linearly proportional to the epistemic uncertainty $U_E(x)$,
 1628 scaled by the temperature τ and shifted by a constant. Therefore, a higher free energy value directly
 1629 corresponds to higher epistemic uncertainty. \square
 1630

1631 G.3 PROOF OF THEOREM 2.2

1633 *Proof.* The proof proceeds by first establishing the relationship between the negative free energy
 1634 $-F(x)$ and the maximum logit, and then arguing that the expected maximum logit decreases as
 1635 sample difficulty increases for a well-trained model.

1636 **Step 1: Relating Negative Free Energy to the Maximum Logit.** The negative free energy, $-F(x)$,
 1637 is the LogSumExp (LSE) of the scaled logits. The LSE function is a smooth approximation of the
 1638 maximum function and is tightly bounded by it. For any vector $\mathbf{z} \in \mathbb{R}^K$, the sum of exponentials
 1639 can be bounded relative to its maximum term, $z_{\max} = \max_k z_k$:

$$1640 \quad e^{z_{\max}} \leq \sum_{k=1}^K e^{z_k} \leq K \cdot e^{z_{\max}}. \quad (30)$$

1644 By taking the logarithm across all parts of the inequality, we obtain the standard bounds for the LSE
 1645 function:
 1646

$$1647 \quad \max_k z_k \leq \log \sum_{k=1}^K e^{z_k} \leq \max_k z_k + \log K. \quad (31)$$

1649 Applying this to our scaled logits, $z_k = f_k(x)/\tau$, and multiplying by τ gives:
 1650

$$1651 \quad \max_k f_k(x) \leq -F(x) \leq \max_k f_k(x) + \tau \log K. \quad (32)$$

1652 This inequality demonstrates that $-F(x)$ is a tight, monotonically increasing function of the maxi-
 1653 mum logit, $\max_k f_k(x)$. Therefore, proving Theorem 2.2 is equivalent to proving that the expected
 1654 maximum logit is a strictly monotonically decreasing function of difficulty:
 1655

$$1656 \quad \mathbb{E}[\max_k f_k(X) \mid D(X, Y_{\text{true}}) = d_1] > \mathbb{E}[\max_k f_k(X) \mid D(X, Y_{\text{true}}) = d_2]. \quad (33)$$

1657 **Step 2: Characterizing the Maximum Logit by Difficulty Level.** We analyze the maximum logit
 1658 for samples conditioned on their difficulty.
 1659

- 1660 • **Low Difficulty ($d = 1$):** A sample (x, y) has difficulty $d = 1$ if and only if its true label
 1661 y receives the highest logit. Thus, for this subpopulation of data, the maximum logit is the
 1662 logit of the true class:
 1663

$$1664 \quad D(x, y_{\text{true}}) = 1 \implies \max_k f_k(x) = f_y(x). \quad (34)$$

1666 A model trained via a standard objective like cross-entropy is explicitly optimized to in-
 1667 crease the value of $f_y(x)$ for all training samples. Consequently, the set of samples where
 1668 the model succeeds ($d = 1$) corresponds to inputs for which the model produces a large,
 1669 dominant logit for the correct class.

- 1670 • **High Difficulty ($d > 1$):** A sample (x, y) has difficulty $d > 1$ if and only if the model's
 1671 prediction is incorrect. This implies that the maximum logit corresponds to an incorrect
 1672 class $k' \neq y$:

$$1673 \quad D(x, y_{\text{true}}) = d > 1 \implies \max_k f_k(x) = f_{k'}(x) \text{ for some } k' \neq y. \quad (35)$$

1674
 1675 **Step 3: Comparing Conditional Expectations.** We compare the expected maximum logit over
 1676 the subpopulation of correctly classified samples ($d_1 = 1$) versus incorrectly classified samples
 1677 ($d_2 > 1$). The training objective directly pushes the values in the set $\{f_Y(X) \mid D(X, Y_{\text{true}}) = 1\}$
 1678 to be as large as possible. In contrast, the values in the set $\{\max_k f_k(X) \mid D(X, Y_{\text{true}}) > 1\}$ arise
 1679 from the model’s failure to generalize.

1680 A fundamental property of a successfully trained and well-generalized, and well-calibrated classifier
 1681 is that its confidence on the examples it classifies correctly is, on average, higher than its confidence
 1682 on the examples it classifies incorrectly (Guo et al., 2017). If this were not the case, the model would
 1683 not have learned a meaningful decision boundary from the data. Thus, the average maximum logit
 1684 for the population of “easy” samples must be greater than that for the population of “hard” samples.
 1685 Formally, for $d_1 < d_2$:

$$\mathbb{E}[\max_k f_k(X) \mid D(X, Y_{\text{true}}) = d_1] > \mathbb{E}[\max_k f_k(X) \mid D(X, Y_{\text{true}}) = d_2]. \quad (36)$$

1686 Given the monotonic relationship established in Equation 32, it follows directly that the expected
 1687 negative free energy also decreases with increasing difficulty. This completes the proof and aligns
 1688 with our empirical observation in Section 2. \square

1691 G.4 PROOF OF PROPOSITION 2.3

1692 *Proof.* The proof follows directly from the definition of a conformal prediction set. By definition,
 1693 the prediction set $\mathcal{C}_G(x)$ for a new instance x includes all labels y for which the scaled nonconformity
 1694 score does not exceed the calibrated quantile $\hat{q}_{1-\alpha}^{(G)}$.

$$\mathcal{C}_G(x) = \left\{ y \mid S_G(x, y) \leq \hat{q}_{1-\alpha}^{(G)} \right\}. \quad (37)$$

1695 Substituting the definition of the scaled score, $S_G(x, y) = G(x)S(x, y)$, we have:

$$\mathcal{C}_G(x) = \left\{ y \mid G(x)S(x, y) \leq \hat{q}_{1-\alpha}^{(G)} \right\}. \quad (38)$$

1696 Since we assume $G(x)$ is strictly positive, we can divide both sides of the inequality by $G(x)$ without
 1697 changing its direction:

$$\mathcal{C}_G(x) = \left\{ y \mid S(x, y) \leq \frac{\hat{q}_{1-\alpha}^{(G)}}{G(x)} \right\}. \quad (39)$$

1698 By defining the instance-adaptive threshold $\theta(x) = \hat{q}_{1-\alpha}^{(G)}/G(x)$, we arrive at the equivalent formu-
 1699 lation:

$$\mathcal{C}_G(x) = \{y \mid S(x, y) \leq \theta(x)\}. \quad (40)$$

1700 This concludes the proof. \square

1701 *Remark G.2.* It is important to emphasize that the new quantile, $\hat{q}_{1-\alpha}^{(G)}$, is fundamentally different
 1702 from the baseline quantile, $\hat{q}_{1-\alpha}$, which would be computed from the unscaled scores.

1703 Let $\mathcal{S}_{\text{base}} = \{S(X_i, Y_i)\}_{i=1}^n$ be the set of baseline calibration scores, and $\mathcal{S}_G =$
 1704 $\{G(X_i)S(X_i, Y_i)\}_{i=1}^n$ be the set of Energy-reweighted calibration scores.

- 1705 • $\hat{q}_{1-\alpha}$ is the $(1 - \alpha)$ -quantile of the empirical distribution defined by $\mathcal{S}_{\text{base}}$.
- 1706 • $\hat{q}_{1-\alpha}^{(G)}$ is the $(1 - \alpha)$ -quantile of the empirical distribution defined by \mathcal{S}_G .

1707 There is no simple, closed-form relationship between $\hat{q}_{1-\alpha}$ and $\hat{q}_{1-\alpha}^{(G)}$. Scaling each score S_i by
 1708 a different factor $G(X_i)$ changes the distribution of scores, including the relative ordering of the
 1709 calibration samples. For instance, a sample (X_j, Y_j) that had a median score in $\mathcal{S}_{\text{base}}$ might have a
 1710 very high score in \mathcal{S}_G if its corresponding energy factor $G(X_j)$ is large.

1711 Consequently, the sample that happens to fall at the $\lceil (1 - \alpha)(n + 1) \rceil$ -th position (thus defining the
 1712 quantile) will almost certainly be different in the baseline and Energy-based cases. In other words,
 1713 one cannot simply take the baseline quantile $\hat{q}_{1-\alpha}$ and scale it by some factor. The entire set of
 1714 calibration scores must be re-computed and re-sorted to find the new, correct quantile $\hat{q}_{1-\alpha}^{(G)}$.

1728 G.5 PROOF OF THEOREM 3.1
1729

1730 *Proof.* The proof relies on the connection between the logits of a classifier trained with cross-entropy
1731 and the Bayesian posterior probability, which shows that the model’s parameters internalize the
1732 training set’s class priors.

1733 **Step 1: Bayesian Decomposition of Logits.** A classifier trained to minimize cross-entropy loss
1734 learns to approximate the posterior probability $P(Y = y|X = x)$. Its logits $f_y(x)$ thus approximate
1735 the log-posterior, up to an instance-specific normalization constant. Using Bayes’ rule, we can
1736 decompose the log-posterior:

$$1737 \log P(Y = y|X = x) = \log P(X = x|Y = y) + \log P(Y = y) - \log P(X = x). \quad (41)$$

1739 When trained on P_{train} , the model’s logits learn to reflect this structure:

$$1740 \quad f_y(x) \approx \log P_{\text{train}}(X = x|Y = y) + \log P_{\text{train}}(Y = y) + C(x), \quad (42)$$

1742 where the term $C(x)$ absorbs instance-dependent factors like $-\log P_{\text{train}}(X = x)$ and other model-
1743 specific biases. Critically, the logit $f_y(x)$ encodes the log-prior probability of class y from the
1744 training distribution.

1745 **Step 2: Connecting Negative Free Energy to the Maximum Logit.** As established in the proof of
1746 Theorem 2.2, the negative free energy $-F(x)$ is tightly and monotonically related to the maximum
1747 logit, $-F(x) \approx \max_k f_k(x)$. For a reasonably accurate classifier, the expectation of the maximum
1748 logit over test samples from a given class y is dominated by instances where the model is correct.
1749 For a correct classification of a sample (x, y) , the maximum logit is the logit of the true class:
1750 $\max_k f_k(x) = f_y(x)$. Building on this, we can state that the expected negative free energy for class
1751 y is primarily driven by the expected logit for that class:

$$1752 \quad \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y)}[-F(X)] \approx \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y)}[f_y(X)]. \quad (43)$$

1754 **Step 3: Comparing Expected Logits for Majority and Minority Classes.** Using the decomposi-
1755 tion from Equation 42, we can express the expected logit for a class y as:

$$1757 \quad \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y)}[f_y(X)] \approx \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y)}[\log P_{\text{train}}(X|Y = y) + C(X)] + \log P_{\text{train}}(Y = y). \quad (44)$$

1758 Let us define the term $A(y) = \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y)}[\log P_{\text{train}}(X|Y = y) + C(X)]$. This term represents
1759 the average “data-fit” or “evidence” for class y , as learned by the model. Under our assumption that
1760 classes y_{maj} and y_{min} have comparable intrinsic complexity and are well-represented, this evidence
1761 term should be similar for both, i.e., $A(y_{\text{maj}}) \approx A(y_{\text{min}})$.

1762 We can now compare the expected negative free energy for the two classes:

$$1764 \quad \mathbb{E}[-F(X)|Y = y_{\text{maj}}] \approx A(y_{\text{maj}}) + \log P_{\text{train}}(Y = y_{\text{maj}}) \quad (45)$$

$$1765 \quad \mathbb{E}[-F(X)|Y = y_{\text{min}}] \approx A(y_{\text{min}}) + \log P_{\text{train}}(Y = y_{\text{min}}) \quad (46)$$

1767 By the proposition’s premise, $P_{\text{train}}(Y = y_{\text{maj}}) > P_{\text{train}}(Y = y_{\text{min}})$, which implies $\log P_{\text{train}}(Y =$
1768 $y_{\text{maj}}) > \log P_{\text{train}}(Y = y_{\text{min}})$. Since $A(y_{\text{maj}}) \approx A(y_{\text{min}})$, the additive log-prior term learned during
1769 training becomes the dominant factor driving the difference. Therefore, we conclude that:

$$1770 \quad \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y_{\text{maj}})}[-F(X)] > \mathbb{E}_{X \sim P_{\text{test}}(X|Y=y_{\text{min}})}[-F(X)]. \quad (47)$$

1772 This result confirms that the model’s systematically higher epistemic uncertainty (lower negative
1773 free energy) for minority classes is a bias inherited from the training distribution’s class priors. This
1774 aligns with prior work showing that models have lower expected logits for minority classes (Ren
1775 et al., 2020; Lyu et al., 2025; Kato & Hotta, 2023; Chen & Su, 2023). \square

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1782 H ANALYSIS OF SOFTMAX SATURATION AND ENERGY-BASED ADAPTIVITY 1783

1784 In this section, we provide a formal mechanism linking the limitations of softmax-based noncon-
1785 formity scores to conformal inefficiency. We demonstrate that while softmax probabilities saturate
1786 rapidly away from the decision boundary, thereby losing information about sample difficulty, the
1787 Helmholtz Free Energy retains this geometric information. We validate this analysis with a toy
1788 experiment visualizing the decision landscapes.

1790 H.1 DISTANCE TO DECISION BOUNDARY AND LOGIT MAGNITUDE 1791

1792 Consider a deep classifier $f : \mathcal{X} \rightarrow \mathbb{R}^K$. For a given input x , let $\hat{y} = \arg \max_k f_k(x)$ be the
1793 predicted class. The decision boundary between the predicted class \hat{y} and the second most likely
1794 class j is defined by the hyperplane where $f_{\hat{y}}(x) = f_j(x)$.

1795 It has been established that for neural networks, the magnitude of the logit vector typically scales
1796 with the distance of the input from the decision boundary (Hein et al., 2019). Let $\delta(x)$ denote the
1797 geometric distance of x from the decision boundary. We observe the following proportionality:

$$1798 \quad 1799 \quad \delta(x) \propto \max_k f_k(x). \quad (48)$$

1800 Therefore, an “easy” sample (one far from the boundary in a high-density region) is characterized by
1801 logits with large magnitudes, while a “hard” sample (near the boundary) yields logits with smaller
1802 or entangled magnitudes. Ideally, an adaptive conformal predictor should produce smaller sets as
1803 $\delta(x)$ increases.

1805 H.2 THE SATURATION OF SOFTMAX AND ENTROPY 1806

1807 Standard conformal scores rely on the softmax distribution $\hat{\pi}(y|x) = \exp(f_y(x)) / \sum_k \exp(f_k(x))$.
1808 A critical limitation of this mapping is *gradient saturation*.

1809 Consider a sample x moving away from the decision boundary such that its logit magnitude scales
1810 by a factor $\alpha > 1$. As $\alpha \rightarrow \infty$, $\hat{\pi}(\hat{y}|x) \rightarrow 1$. The gradient of the softmax output with respect to the
1811 dominant logit $f_{\hat{y}}$ is given by:

$$1812 \quad 1813 \quad \frac{\partial \hat{\pi}(\hat{y}|x)}{\partial f_{\hat{y}}} = \hat{\pi}(\hat{y}|x)(1 - \hat{\pi}(\hat{y}|x)). \quad (49)$$

1815 As $\hat{\pi} \rightarrow 1$, this gradient approaches 0. Similarly, the Shannon Entropy $\mathbb{H}(x)$ of the distribution
1816 approaches 0.

1818 **Implication for Conformal Prediction:** Once a sample is sufficiently far from the boundary to
1819 achieve a high confidence (e.g., $\hat{\pi} > 0.99$), the softmax score saturates. The model becomes geo-
1820 metrically insensitive: a sample at distance d and a sample at distance $10d$ yield indistinguishable
1821 conformal scores. This saturation restricts the adaptive capacity of the prediction sets. For high-
1822 confidence samples, the sets cannot achieve greater efficiency because the score yields no further
1823 signal.

1824 H.3 NON-SATURATION OF FREE ENERGY 1825

1826 In contrast, the negative Helmholtz Free Energy is defined as $-F(x) = \tau \log \sum_k \exp(f_k(x)/\tau)$.
1827 As derived in Appendix G.3, this quantity is bounded by the maximum logit:

$$1828 \quad 1829 \quad -F(x) \approx \max_k f_k(x). \quad (50)$$

1830 Unlike softmax, the Free Energy does not saturate. Its derivative with respect to the dominant logit
1831 is approximately 1:

$$1832 \quad 1833 \quad \frac{\partial(-F(x))}{\partial f_{\hat{y}}} \approx 1. \quad (51)$$

1835 This indicates that $-F(x)$ grows linearly with the logit magnitude, thereby acting as a faithful proxy
for the distance $\delta(x)$ even in high-confidence regimes.

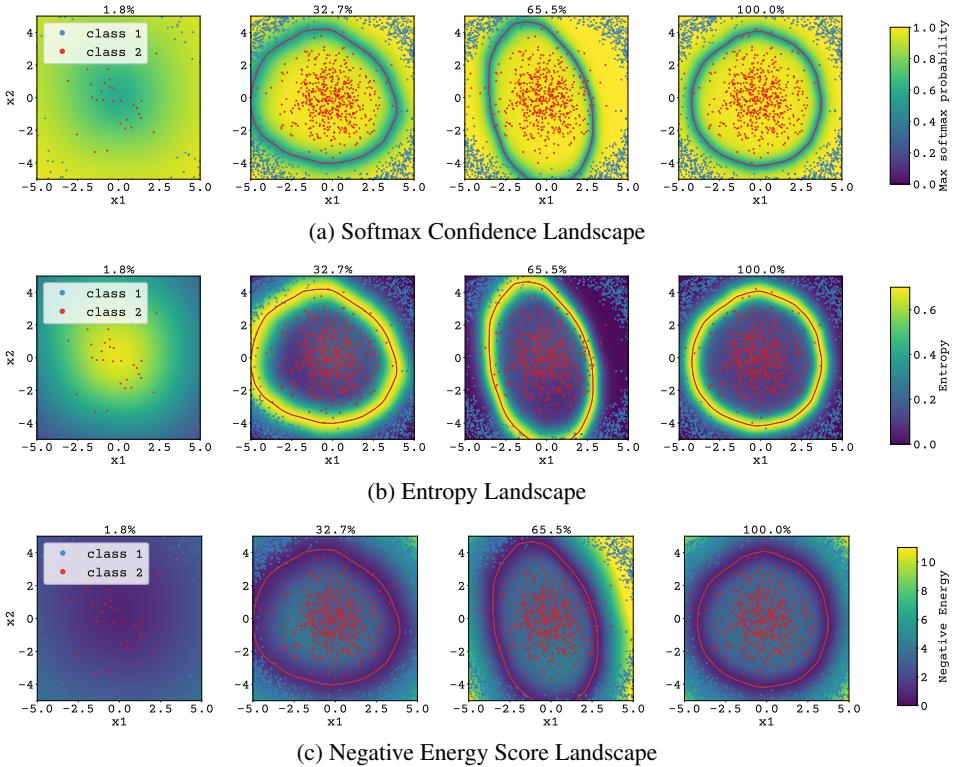
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H.4 EMPIRICAL VISUALIZATION OF UNCERTAINTY LANDSCAPES

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To empirically validate this behavior, we trained a 3-layer Multilayer Perceptron (MLP) on a 2D toy dataset consisting of two concentric classes. Figure 6 visualizes the value of Max Softmax Confidence, Shannon Entropy, and Negative Free Energy across the input space \mathcal{X} .

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(a) Softmax Confidence Landscape

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(b) Entropy Landscape

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(c) Negative Energy Score Landscape

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Figure 6: Evolution of uncertainty landscapes on a 2D toy dataset throughout the training process. Columns represent progressive checkpoints from the early training phase (left) to full convergence (right). The red line indicates the decision boundary. (a) Softmax probabilities saturate rapidly. The yellow region (confidence ≈ 1.0) is flat, making points near the boundary indistinguishable from points far away. (b) Entropy exhibits similar saturation (dark blue region), vanishing to zero for most of the domain. (c) Negative Free Energy retains gradients throughout the domain. Note the continuous color transition scaling with the distance from the decision boundary, identifying “easier” points with higher values even when softmax is saturated.

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H.5 MECHANISM OF ENERGY-BASED EFFICIENCY

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Our proposed method leverages this non-saturating property by modulating the base nonconformity score $S(x, y)$ with a sample-specific scaler $G(x) \propto \text{softplus}(-F(x))$.

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For “easy” samples (large $\delta(x)$), $-F(x)$ is large positive. This results in a large scaling factor $G(x) \gg 1$.

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1. The nonconformity scores for incorrect labels (which are naturally non-zero) are magnified significantly by $G(x)$, pushing them well above the calibrated quantile \hat{q} .
2. The nonconformity score for the true label (typically near zero) remains small even after scaling.

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This amplification forces the exclusion of incorrect classes that might otherwise have been included due to the looseness of the global quantile \hat{q} , thereby reducing the prediction set size. Because $F(x)$ does not saturate, this efficiency gain continues to improve as samples get “easier,” a property unattainable with softmax-based modulation.

1890 I PERFORMANCE ANALYSIS STRATIFIED BY SAMPLE DIFFICULTY

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 1892 In this section, we report the performance of different nonconformity scores by stratifying test sam-
 1893 ples based on their difficulty. Sample difficulty is defined as the rank of the true label in the model’s
 1894 predicted probability ordering; a lower rank indicates an “easier” sample, while a higher rank signi-
 1895 fies a “harder”, often misclassified, one.

1896 In Section 3.1, we established that energy-reweighting reduces the overall average prediction set
 1897 size while maintaining target coverage. However, this aggregate metric does not reveal whether
 1898 these efficiency gains are distributed evenly or are concentrated on specific types of samples. This
 1899 stratified analysis provides a more granular view to answer that question.

1900 We expect an adaptive method to produce the most significant set size reductions for easy samples,
 1901 where the model is confident, while appropriately adjusting for harder samples where uncertainty
 1902 is higher. Table 6 presents these stratified results, detailing how coverage and average set size vary
 1903 across the different difficulty levels.

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 1905 Table 6: Coverage and average set size on ImageNet, stratified by sample difficulty. Results are for a ResNet-
 1906 50 model at $\alpha = 0.01$ and averaged over 10 trials. The table compares baseline adaptive scores with their
 1907 energy-based variants, which generally produce smaller sets for easier samples while maintaining coverage.

Difficulty	Count	APS				RAPS				SAPS			
		w/o Energy		w/ Energy		w/o Energy		w/ Energy		w/o Energy		w/ Energy	
		Cov.	Set Size	Cov.	Set Size	Cov.	Set Size	Cov.	Set Size	Cov.	Set Size	Cov.	Set Size
1 to 1	15990	1.00	39.08	1.00	32.31	1.00	36.97	1.00	30.53	1.00	35.69	1.00	29.11
2 to 3	2547	1.00	38.38	1.00	34.96	1.00	37.17	1.00	34.10	1.00	36.98	1.00	33.69
4 to 6	638	1.00	37.86	1.00	36.46	1.00	37.32	1.00	36.29	1.00	37.35	1.00	36.02
7 to 10	306	1.00	37.66	1.00	37.09	1.00	37.40	1.00	37.20	1.00	37.47	1.00	36.96
11 to 100	453	0.76	37.49	0.76	37.84	0.76	37.51	0.76	38.27	0.76	37.58	0.76	37.98
101 to 1000	66	0.00	37.82	0.00	38.45	0.00	37.60	0.00	38.60	0.00	37.34	0.00	38.05

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J ENERGY-BASED LAC

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The standard LAC score is inversely proportional to the softmax probability. For its energy-based variant, as mentioned in Appendix E.2, we use a base score of $S_{\text{LAC}}(x, y) = -\hat{\pi}(y|x)$. For a difficult input, which corresponds to a low negative free energy score, the objective is to produce a larger prediction set. This requires the nonconformity scores of more classes to fall below the calibrated threshold. Conversely, for an easy input (high negative free energy), the nonconformity scores should be scaled to produce a smaller set. Therefore, for the energy-based LAC, we divide the base score by the energy-based scaling factor. This adjustment ensures that for difficult inputs, the small scaling factor in the denominator makes the nonconformity scores smaller, including more labels in the set.

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So formally, we define the Energy-based LAC nonconformity score as:

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$$S_{\text{EB-LAC}}(x, y) = \frac{-\hat{\pi}(y|x)}{\frac{1}{\beta} \log(1 + e^{-\beta F(x)})} \quad (52)$$

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where $\hat{\pi}(y|x)$ is the softmax probability, $F(x)$ is the Helmholtz free energy, and β is the softplus sharpness parameter.

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J.1 ENERGY-BASED LAC PERFORMANCE IN STANDARD SCENARIO

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Table 7: Performance of the LAC nonconformity score function and its energy-based variant on CIFAR-100, ImageNet, and Places365 at miscoverage levels $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$. Results are averaged over 10 trials. For the **Set Size** column, lower is better. **Bold** values indicate the best performance within the method family (with and without Energy).

Method	$\alpha = 0.1$		$\alpha = 0.05$		$\alpha = 0.025$		$\alpha = 0.01$	
	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size
CIFAR-100 (ResNet-56)								
LAC	w/o Energy	0.90 ± 0.01	2.54 ± 0.06	0.95 ± 0.00	5.26 ± 0.17	0.974 ± 0.001	9.50 ± 0.21	0.99 ± 0.00
	w/ Energy	0.90 ± 0.01	2.52 ± 0.04	0.95 ± 0.00	5.21 ± 0.15	0.974 ± 0.002	9.09 ± 0.16	0.99 ± 0.00
ImageNet (ResNet-50)								
LAC	w/o Energy	0.90 ± 0.00	1.49 ± 0.01	0.95 ± 0.00	2.68 ± 0.05	0.975 ± 0.001	5.48 ± 0.18	0.99 ± 0.00
	w/ Energy	0.90 ± 0.00	1.48 ± 0.01	0.95 ± 0.00	2.68 ± 0.04	0.975 ± 0.001	5.42 ± 0.15	0.99 ± 0.00
Places365 (ResNet-50)								
LAC	w/o Energy	0.90 ± 0.001	6.21 ± 0.03	0.95 ± 0.00	11.36 ± 0.04	0.973 ± 0.001	19.55 ± 0.41	0.99 ± 0.001
	w/ Energy	0.90 ± 0.001	6.19 ± 0.03	0.95 ± 0.001	11.11 ± 0.11	0.973 ± 0.001	19.28 ± 0.43	0.99 ± 0.001

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J.2 ENERGY-BASED LAC PERFORMANCE WITH IMBALANCED TRAINING PRIORS

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2003Table 8: Performance of LAC and its energy-based variant on imbalanced CIFAR-100 with varying imbalance factors ($\lambda \in \{0.005, 0.01, 0.02, 0.03\}$) and at miscoverage levels $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$. Results are averaged over 10 trials with a ResNet-56 model. For the average set size, lower is better. **Bold** values indicate the best performance.

Method	$\alpha = 0.1$		$\alpha = 0.05$		$\alpha = 0.025$		$\alpha = 0.01$	
	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size
CIFAR-100-LT ($\lambda = 0.005$, mild imbalance) (ResNet-56)								
LAC	w/o Energy	0.897 ± 0.007	7.04 ± 0.24	0.947 ± 0.004	12.98 ± 0.49	0.973 ± 0.003	21.31 ± 0.84	0.988 ± 0.002
	w/ Energy	0.897 ± 0.006	6.91 ± 0.19	0.947 ± 0.005	12.68 ± 0.51	0.973 ± 0.004	20.56 ± 1.01	0.989 ± 0.002
CIFAR-100-LT ($\lambda = 0.01$) (ResNet-56)								
LAC	w/o Energy	0.900 ± 0.007	11.92 ± 0.41	0.951 ± 0.003	20.75 ± 0.51	0.975 ± 0.003	30.58 ± 0.68	0.990 ± 0.001
	w/ Energy	0.900 ± 0.007	11.52 ± 0.32	0.950 ± 0.003	20.32 ± 0.39	0.976 ± 0.003	30.51 ± 0.64	0.990 ± 0.001
CIFAR-100-LT ($\lambda = 0.02$) (ResNet-56)								
LAC	w/o Energy	0.901 ± 0.007	27.78 ± 0.73	0.950 ± 0.007	42.03 ± 1.67	0.975 ± 0.003	54.88 ± 1.13	0.990 ± 0.002
	w/ Energy	0.900 ± 0.007	27.63 ± 0.65	0.951 ± 0.006	41.49 ± 1.19	0.976 ± 0.003	54.13 ± 0.93	0.990 ± 0.001
CIFAR-100-LT ($\lambda = 0.03$, severe imbalance) (ResNet-56)								
LAC	w/o Energy	0.901 ± 0.006	28.34 ± 0.47	0.951 ± 0.004	41.79 ± 0.73	0.976 ± 0.002	54.71 ± 0.82	0.990 ± 0.002
	w/ Energy	0.901 ± 0.006	27.93 ± 0.42	0.952 ± 0.004	41.50 ± 0.71	0.975 ± 0.003	54.02 ± 1.07	0.990 ± 0.003

J.3 ENERGY-BASED LAC PERFORMANCE ANALYSIS STRATIFIED BY SAMPLE DIFFICULTY

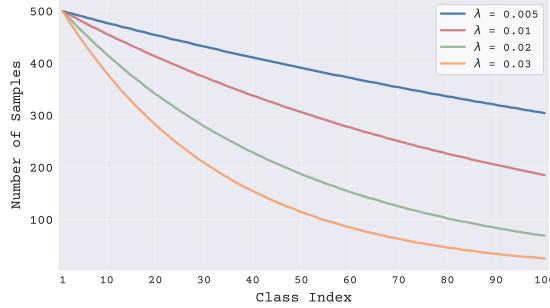
2025
2026Table 9: Coverage and average set size for the LAC method on ImageNet, stratified by sample difficulty. Results are shown for $\alpha = 0.01$ and $\alpha = 0.025$. The table compares the baseline LAC with its energy-based variant.

Difficulty Level	Count	LAC							
		$\alpha = 0.01$				$\alpha = 0.025$			
		w/o Energy		w/ Energy		w/o Energy		w/ Energy	
		Cov.	Set Size	Cov.	Set Size	Cov.	Set Size	Cov.	Set Size
1 to 1	15990	1.00	10.77	1.00	9.24	1.00	4.34	1.00	4.25
2 to 3	2547	1.00	21.47	1.00	20.61	0.99	8.86	0.99	8.78
4 to 6	638	0.99	32.07	0.98	32.96	0.93	13.29	0.93	13.36
7 to 10	306	0.96	38.18	0.96	40.00	0.84	15.86	0.85	16.14
11 to 25	275	0.92	44.22	0.91	47.14	0.60	17.07	0.56	17.05
26 to 50	104	0.62	47.37	0.61	50.88	0.04	17.60	0.06	17.59
51 to 100	74	0.32	56.19	0.39	64.61	0.00	19.99	0.00	21.07

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2052 K ADDITIONAL EXPERIMENTAL RESULTS FOR IMBALANCED DATA

2054 To evaluate performance under class imbalance as described in Section 3.2, we construct several
 2055 long-tailed variants of the CIFAR-100 dataset. In these variants, the number of training samples per
 2056 class follows an exponential decay controlled by an imbalance factor λ . Figure 7 illustrates how
 2057 different values of λ create varying levels of imbalance in the training distribution.



2069 Figure 7: Class distributions under varying imbalance levels. The number of samples per class follows an
 2070 exponential decay pattern proportional to $\exp(-\lambda \cdot j)$, where larger λ values induce stronger imbalance.
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2072 K.1 RESULTS FOR DIFFERENT IMBALANCE FACTOR λ

2074 Table 10: Performance comparison of APS, RAPS, SAPS, and their Energy-based variants on imbal-
 2075 anced CIFAR-100 with varying imbalance factors ($\lambda \in \{0.01, 0.02, 0.03\}$) and at miscoverage levels
 2076 $\alpha \in \{0.01, 0.025, 0.05, 0.1\}$. Results are averaged over 10 trials with a ResNet-56 model. For the **Set Size**
 2077 column, lower is better. **Bold** values indicate the best performance within each method family.

Method	$\alpha = 0.1$		$\alpha = 0.05$		$\alpha = 0.025$		$\alpha = 0.01$	
	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size	Coverage	Set Size
CIFAR-100-LT ($\lambda = 0.01$, mild imbalance) (ResNet-56)								
APS	w/o Energy	0.90 \pm 0.01	14.56 \pm 0.46	0.95 \pm 0.00	25.59 \pm 0.25	0.975 \pm 0.003	37.56 \pm 0.98	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	11.86 \pm 0.37	0.95 \pm 0.00	21.44 \pm 0.52	0.975 \pm 0.003	31.70 \pm 0.97	0.99 \pm 0.00
RAPS	w/o Energy	0.90 \pm 0.01	15.48 \pm 0.47	0.95 \pm 0.00	27.62 \pm 0.33	0.974 \pm 0.003	40.87 \pm 1.00	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	11.77 \pm 0.42	0.95 \pm 0.00	21.57 \pm 0.56	0.975 \pm 0.003	31.93 \pm 0.98	0.99 \pm 0.00
SAPS	w/o Energy	0.90 \pm 0.01	15.02 \pm 0.50	0.95 \pm 0.00	26.98 \pm 0.32	0.974 \pm 0.002	40.38 \pm 1.26	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	11.80 \pm 0.40	0.95 \pm 0.00	21.35 \pm 0.55	0.975 \pm 0.003	31.84 \pm 1.02	0.99 \pm 0.00
CIFAR-100-LT ($\lambda = 0.02$) (ResNet-56)								
APS	w/o Energy	0.90 \pm 0.01	29.62 \pm 0.56	0.95 \pm 0.01	45.76 \pm 1.32	0.975 \pm 0.003	59.26 \pm 0.82	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	28.23 \pm 0.74	0.95 \pm 0.01	42.36 \pm 1.13	0.976 \pm 0.003	54.95 \pm 0.91	0.99 \pm 0.00
RAPS	w/o Energy	0.90 \pm 0.01	32.86 \pm 0.81	0.95 \pm 0.01	52.01 \pm 1.55	0.976 \pm 0.003	66.83 \pm 1.18	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	28.72 \pm 0.77	0.95 \pm 0.01	42.72 \pm 1.02	0.976 \pm 0.003	56.31 \pm 0.86	0.99 \pm 0.00
SAPS	w/o Energy	0.90 \pm 0.01	32.37 \pm 0.83	0.95 \pm 0.01	51.47 \pm 1.48	0.976 \pm 0.003	66.28 \pm 1.21	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	28.73 \pm 0.78	0.95 \pm 0.01	42.63 \pm 1.09	0.976 \pm 0.003	56.16 \pm 0.89	0.99 \pm 0.00
CIFAR-100-LT ($\lambda = 0.03$, severe imbalance) (ResNet-56)								
APS	w/o Energy	0.90 \pm 0.01	30.35 \pm 0.57	0.95 \pm 0.00	44.45 \pm 0.79	0.975 \pm 0.002	58.05 \pm 0.95	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.00	28.42 \pm 0.52	0.95 \pm 0.00	42.40 \pm 0.63	0.975 \pm 0.003	55.61 \pm 1.54	0.99 \pm 0.00
RAPS	w/o Energy	0.90 \pm 0.01	34.47 \pm 0.72	0.95 \pm 0.01	49.94 \pm 1.14	0.975 \pm 0.003	64.66 \pm 1.13	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	29.29 \pm 0.66	0.95 \pm 0.00	43.68 \pm 0.78	0.975 \pm 0.003	57.23 \pm 1.02	0.99 \pm 0.00
SAPS	w/o Energy	0.90 \pm 0.01	34.05 \pm 0.71	0.95 \pm 0.01	49.45 \pm 1.33	0.974 \pm 0.003	64.35 \pm 1.33	0.99 \pm 0.00
	w/ Energy	0.90 \pm 0.01	29.19 \pm 0.61	0.95 \pm 0.01	43.61 \pm 0.85	0.975 \pm 0.003	57.12 \pm 0.97	0.99 \pm 0.00

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K.2 PERFORMANCE UNDER IMBALANCED CALIBRATION AND TEST SETS

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The results reported in Table 2 and Table 10 evaluate models trained on imbalanced data but calibrated and tested on balanced sets. We now consider a more realistic scenario where the calibration and test sets also follow the same imbalanced distribution as the training set.

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This setting is particularly challenging for standard CP methods, as the limited number of calibration samples for minority classes can impede reliable coverage guarantees. In such cases, approaches like clustered conformal prediction Ding et al. (2023b), which can provide coverage with fewer calibration samples, are practical alternatives.

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Our energy-based method is designed to address this challenge by adaptively enlarging prediction sets for uncertain inputs, which often correspond to minority class samples. To benchmark this behavior, we compare it against another principled reweighting strategy that directly uses class priors. This approach was introduced by Ding et al. (2025), who proposed the **Prevalence-Adjusted Softmax (PAS)** score. The PAS score modifies the non-adaptive score by dividing the negative softmax probability (LACscore without bias term) by the empirical class prior, $\hat{p}(y)$, to improve coverage for rare classes. The nonconformity score is defined as:

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$$S_{\text{PAS}}(x, y) = \frac{-\hat{\pi}(y|x)}{\hat{p}(y)} \quad (53)$$

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We extend this concept to adaptive nonconformity scores such as APS and RAPS. For these scores, a smaller value indicates higher conformity. To increase the likelihood of including labels from rare classes (which have a small $\hat{p}(y)$), we multiply the base score by the class prior. This makes the minority classes more likely to be included in the final prediction set. We refer to this method as **Prevalence-Adjusted (PA) Nonconformity Scores**. The general formulation is:

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$$S_{\text{PA}}(x, y) = S_{\text{adaptive}}(x, y) \cdot \hat{p}(y) \quad (54)$$

where $S_{\text{adaptive}}(x, y)$ is an adaptive score like $S_{\text{APS}}(x, y)$.

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Given this, Table 11 and Table 12 present a comparison between the standard adaptive scores, our energy-based variants, and the prevalence-adjusted variants in this fully imbalanced setting. We report marginal coverage, average set size, and MacroCov to provide a comprehensive view of performance.

Table 11: Performance on fully imbalanced CIFAR-100-LT for high confidence levels ($\alpha \in \{0.025, 0.01\}$). For each method, we compare the **Standard** baseline, the **Prevalence-Adj.** variant, and our **Energy-based** variant. Lower **Set Size** is better.

Method	Variant	$\alpha = 0.025$			$\alpha = 0.01$		
		Cov	Size	MacroCov	Cov	Size	MacroCov
CIFAR-100-LT ($\lambda = 0.005$, mild imbalance)							
LAC	Standard	0.97 ± 0.00	21.48 ± 0.86	0.97 ± 0.00	0.99 ± 0.00	34.60 ± 1.77	0.99 ± 0.00
	Prevalence-Adj. (PAS)	0.98 ± 0.00	22.84 ± 0.99	0.98 ± 0.00	0.99 ± 0.00	36.23 ± 0.98	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	21.11 ± 0.79	0.97 ± 0.00	0.99 ± 0.00	33.18 ± 1.44	0.99 ± 0.00
APS	Standard	0.97 ± 0.00	28.61 ± 1.23	0.97 ± 0.00	0.99 ± 0.00	46.37 ± 1.39	0.99 ± 0.00
	Prevalence-Adj.	0.98 ± 0.00	29.72 ± 1.80	0.97 ± 0.00	0.99 ± 0.00	50.57 ± 2.08	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	22.09 ± 1.19	0.97 ± 0.00	0.99 ± 0.00	35.78 ± 1.43	0.99 ± 0.00
RAPS	Standard	0.97 ± 0.00	30.52 ± 1.44	0.97 ± 0.00	0.99 ± 0.00	51.46 ± 1.84	0.99 ± 0.00
	Prevalence-Adj.	0.97 ± 0.00	31.45 ± 1.66	0.97 ± 0.00	0.99 ± 0.00	51.41 ± 1.66	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	22.57 ± 1.08	0.97 ± 0.00	0.99 ± 0.00	36.75 ± 1.42	0.99 ± 0.00
SAPS	Standard	0.97 ± 0.00	29.54 ± 1.48	0.97 ± 0.00	0.99 ± 0.00	50.35 ± 2.04	0.99 ± 0.00
	Prevalence-Adj.	0.97 ± 0.00	30.41 ± 1.62	0.97 ± 0.00	0.99 ± 0.00	50.50 ± 1.33	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	22.36 ± 1.13	0.97 ± 0.00	0.99 ± 0.00	36.39 ± 1.23	0.99 ± 0.00
CIFAR-100-LT ($\lambda = 0.01$)							
LAC	Standard	0.97 ± 0.00	30.22 ± 0.83	0.97 ± 0.00	0.99 ± 0.00	45.45 ± 1.77	0.99 ± 0.00
	Prevalence-Adj. (PAS)	0.98 ± 0.00	33.95 ± 1.00	0.98 ± 0.00	0.99 ± 0.00	50.56 ± 1.59	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	29.93 ± 0.99	0.97 ± 0.00	0.99 ± 0.00	43.94 ± 1.34	0.99 ± 0.00
APS	Standard	0.97 ± 0.00	36.97 ± 1.61	0.97 ± 0.00	0.99 ± 0.00	54.90 ± 2.32	0.99 ± 0.00
	Prevalence-Adj.	0.98 ± 0.00	43.29 ± 2.15	0.98 ± 0.00	0.99 ± 0.00	66.30 ± 2.14	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	31.20 ± 1.12	0.97 ± 0.00	0.99 ± 0.00	45.15 ± 1.81	0.99 ± 0.00
RAPS	Standard	0.97 ± 0.00	40.39 ± 1.29	0.97 ± 0.00	0.99 ± 0.00	60.09 ± 2.93	0.99 ± 0.00
	Prevalence-Adj.	0.98 ± 0.00	43.21 ± 2.15	0.98 ± 0.00	0.99 ± 0.00	65.82 ± 2.35	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	31.64 ± 1.24	0.97 ± 0.00	0.99 ± 0.00	45.99 ± 1.67	0.99 ± 0.00
SAPS	Standard	0.97 ± 0.00	39.81 ± 1.53	0.97 ± 0.00	0.99 ± 0.00	59.11 ± 2.61	0.99 ± 0.00
	Prevalence-Adj.	0.98 ± 0.00	42.34 ± 2.12	0.98 ± 0.00	0.99 ± 0.00	65.18 ± 2.45	0.99 ± 0.00
	Energy-based	0.97 ± 0.01	31.46 ± 1.26	0.97 ± 0.01	0.99 ± 0.00	45.73 ± 1.71	0.99 ± 0.00
CIFAR-100-LT ($\lambda = 0.02$)							
LAC	Standard	0.97 ± 0.00	52.86 ± 1.54	0.97 ± 0.00	0.99 ± 0.00	66.53 ± 1.36	0.99 ± 0.00
	Prevalence-Adj. (PAS)	0.98 ± 0.00	61.74 ± 1.37	0.98 ± 0.00	0.99 ± 0.00	75.07 ± 1.20	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	51.96 ± 1.30	0.97 ± 0.00	0.99 ± 0.00	65.48 ± 2.27	0.99 ± 0.00
APS	Standard	0.96 ± 0.00	56.87 ± 1.30	0.97 ± 0.00	0.99 ± 0.00	71.19 ± 0.92	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	74.42 ± 0.39	0.98 ± 0.00	1.00 ± 0.00	86.18 ± 0.30	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	53.13 ± 1.53	0.97 ± 0.00	0.99 ± 0.00	68.02 ± 1.60	0.99 ± 0.00
RAPS	Standard	0.96 ± 0.00	61.92 ± 1.24	0.97 ± 0.00	0.99 ± 0.00	76.89 ± 1.40	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	73.38 ± 0.28	0.98 ± 0.00	1.00 ± 0.00	86.52 ± 0.40	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	54.13 ± 1.37	0.97 ± 0.00	0.99 ± 0.00	69.14 ± 1.92	0.99 ± 0.00
SAPS	Standard	0.96 ± 0.00	61.42 ± 1.18	0.97 ± 0.00	0.99 ± 0.00	76.01 ± 1.11	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	73.12 ± 0.36	0.98 ± 0.00	1.00 ± 0.00	86.15 ± 0.29	0.99 ± 0.00
	Energy-based	0.97 ± 0.00	54.02 ± 1.58	0.97 ± 0.00	0.99 ± 0.00	69.01 ± 1.67	0.99 ± 0.00
CIFAR-100-LT ($\lambda = 0.03$, severe imbalance)							
LAC	Standard	0.96 ± 0.00	51.39 ± 1.25	0.97 ± 0.00	0.98 ± 0.00	67.76 ± 0.84	0.99 ± 0.00
	Prevalence-Adj. (PAS)	0.98 ± 0.00	62.93 ± 0.99	0.98 ± 0.00	1.00 ± 0.00	75.03 ± 0.46	0.99 ± 0.00
	Energy-based	0.96 ± 0.00	50.77 ± 1.35	0.97 ± 0.00	0.98 ± 0.00	66.87 ± 0.95	0.99 ± 0.00
APS	Standard	0.96 ± 0.00	53.95 ± 1.21	0.97 ± 0.00	0.99 ± 0.00	69.98 ± 1.29	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	79.13 ± 0.69	0.98 ± 0.00	1.00 ± 0.00	86.14 ± 0.28	0.99 ± 0.00
	Energy-based	0.96 ± 0.01	52.00 ± 1.48	0.97 ± 0.00	0.99 ± 0.00	69.77 ± 1.34	0.99 ± 0.00
RAPS	Standard	0.96 ± 0.00	60.09 ± 1.28	0.97 ± 0.00	0.98 ± 0.00	75.77 ± 1.34	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	78.45 ± 0.47	0.98 ± 0.00	1.00 ± 0.00	86.09 ± 0.42	0.99 ± 0.00
	Energy-based	0.96 ± 0.01	53.67 ± 1.54	0.97 ± 0.00	0.99 ± 0.00	71.64 ± 1.17	0.99 ± 0.00
SAPS	Standard	0.96 ± 0.00	59.61 ± 1.28	0.97 ± 0.00	0.98 ± 0.00	75.59 ± 1.15	0.99 ± 0.00
	Prevalence-Adj.	0.99 ± 0.00	78.38 ± 0.51	0.98 ± 0.00	1.00 ± 0.00	86.02 ± 0.33	0.99 ± 0.00
	Energy-based	0.96 ± 0.00	53.68 ± 1.51	0.97 ± 0.00	0.98 ± 0.00	71.72 ± 1.30	0.99 ± 0.00

2214 Table 12: Performance on fully imbalanced CIFAR-100-LT for lower confidence levels ($\alpha \in \{0.1, 0.05\}$). For
 2215 each method, we compare the **Standard** baseline, the **Prevalence-Adj.** variant, and our **Energy-based** variant.
 2216 **Lower Set Size** is better.

2218 2219 2220 2221 2222 2223 2224 2225 2226 2227 2228 2229 2230 2231 2232 2233 2234 2235 2236 2237 2238 2239 2240 2241 2242 2243 2244 2245 2246 2247 2248 2249 2250 2251 2252 2253 2254 2255 2256 2257 2258 2259 2260 2261 2262 2263 2264 2265 2266 2267	Method	Variant	$\alpha = 0.1$			$\alpha = 0.05$		
			Cov	Size	MacroCov	Cov	Size	MacroCov
CIFAR-100-LT ($\lambda = 0.005$, mild imbalance)								
2221 2222 2223	LAC	Standard	0.90 \pm 0.01	7.12 \pm 0.26	0.90 \pm 0.01	0.95 \pm 0.01	13.08 \pm 0.55	0.95 \pm 0.01
		Prevalence-Adj. (PAS)	0.90 \pm 0.01	7.26 \pm 0.30	0.90 \pm 0.01	0.95 \pm 0.01	13.39 \pm 0.60	0.95 \pm 0.01
		Energy-based	0.90 \pm 0.01	6.96 \pm 0.23	0.90 \pm 0.01	0.95 \pm 0.01	12.74 \pm 0.54	0.95 \pm 0.01
2224 2225 2226	APS	Standard	0.90 \pm 0.01	8.48 \pm 0.42	0.90 \pm 0.01	0.95 \pm 0.01	17.55 \pm 1.03	0.95 \pm 0.01
		Prevalence-Adj.	0.90 \pm 0.01	8.81 \pm 0.35	0.90 \pm 0.01	0.95 \pm 0.01	17.95 \pm 0.82	0.95 \pm 0.01
		Energy-based	0.90 \pm 0.01	7.39 \pm 0.22	0.90 \pm 0.01	0.95 \pm 0.01	13.42 \pm 0.56	0.95 \pm 0.01
2227 2228 2229	RAPS	Standard	0.90 \pm 0.01	8.95 \pm 0.44	0.90 \pm 0.01	0.95 \pm 0.01	18.89 \pm 1.06	0.95 \pm 0.01
		Prevalence-Adj.	0.90 \pm 0.01	9.14 \pm 0.46	0.90 \pm 0.01	0.95 \pm 0.01	19.19 \pm 0.89	0.95 \pm 0.01
		Energy-based	0.90 \pm 0.01	7.58 \pm 0.25	0.90 \pm 0.01	0.95 \pm 0.01	13.33 \pm 0.56	0.95 \pm 0.01
2230 2231 2232	SAPS	Standard	0.90 \pm 0.01	8.63 \pm 0.47	0.90 \pm 0.01	0.95 \pm 0.01	18.22 \pm 1.09	0.95 \pm 0.01
		Prevalence-Adj.	0.90 \pm 0.01	8.83 \pm 0.46	0.90 \pm 0.01	0.95 \pm 0.01	18.50 \pm 0.91	0.95 \pm 0.01
		Energy-based	0.90 \pm 0.01	7.53 \pm 0.28	0.90 \pm 0.01	0.95 \pm 0.00	13.30 \pm 0.52	0.95 \pm 0.00
CIFAR-100-LT ($\lambda = 0.01$)								
2233 2234 2235	LAC	Standard	0.89 \pm 0.01	11.37 \pm 0.28	0.89 \pm 0.01	0.95 \pm 0.01	20.09 \pm 0.71	0.95 \pm 0.01
		Prevalence-Adj. (PAS)	0.91 \pm 0.01	12.20 \pm 0.24	0.90 \pm 0.01	0.96 \pm 0.00	22.03 \pm 0.63	0.95 \pm 0.00
		Energy-based	0.89 \pm 0.01	11.12 \pm 0.29	0.89 \pm 0.01	0.95 \pm 0.01	19.73 \pm 0.70	0.95 \pm 0.01
2236 2237 2238	APS	Standard	0.89 \pm 0.00	13.76 \pm 0.30	0.90 \pm 0.00	0.95 \pm 0.01	25.34 \pm 0.95	0.95 \pm 0.01
		Prevalence-Adj.	0.91 \pm 0.01	14.68 \pm 0.36	0.90 \pm 0.01	0.96 \pm 0.01	26.62 \pm 0.96	0.95 \pm 0.01
		Energy-based	0.89 \pm 0.00	11.29 \pm 0.22	0.89 \pm 0.00	0.95 \pm 0.01	20.80 \pm 0.70	0.95 \pm 0.01
2239 2240 2241	RAPS	Standard	0.89 \pm 0.00	14.72 \pm 0.19	0.90 \pm 0.00	0.95 \pm 0.01	26.91 \pm 1.03	0.95 \pm 0.01
		Prevalence-Adj.	0.90 \pm 0.01	15.03 \pm 0.54	0.90 \pm 0.01	0.95 \pm 0.00	27.14 \pm 0.84	0.95 \pm 0.00
		Energy-based	0.89 \pm 0.00	11.22 \pm 0.18	0.89 \pm 0.00	0.95 \pm 0.01	20.95 \pm 0.77	0.95 \pm 0.01
2242 2243 2244	SAPS	Standard	0.89 \pm 0.00	14.22 \pm 0.19	0.90 \pm 0.00	0.95 \pm 0.01	26.19 \pm 0.98	0.95 \pm 0.01
		Prevalence-Adj.	0.90 \pm 0.01	14.55 \pm 0.43	0.90 \pm 0.01	0.95 \pm 0.00	26.65 \pm 0.84	0.95 \pm 0.00
		Energy-based	0.89 \pm 0.00	11.26 \pm 0.17	0.89 \pm 0.00	0.95 \pm 0.01	20.73 \pm 0.80	0.95 \pm 0.01
CIFAR-100-LT ($\lambda = 0.02$)								
2245 2246 2247	LAC	Standard	0.86 \pm 0.01	24.97 \pm 0.67	0.88 \pm 0.01	0.93 \pm 0.01	38.81 \pm 1.07	0.94 \pm 0.01
		Prevalence-Adj. (PAS)	0.91 \pm 0.00	30.25 \pm 0.46	0.91 \pm 0.00	0.96 \pm 0.00	46.98 \pm 1.15	0.96 \pm 0.00
		Energy-based	0.86 \pm 0.01	24.65 \pm 0.67	0.88 \pm 0.01	0.93 \pm 0.01	38.46 \pm 0.94	0.94 \pm 0.01
2248 2249 2250	APS	Standard	0.87 \pm 0.01	26.69 \pm 0.71	0.88 \pm 0.01	0.93 \pm 0.01	42.41 \pm 1.32	0.94 \pm 0.01
		Prevalence-Adj.	0.92 \pm 0.00	37.05 \pm 0.63	0.91 \pm 0.00	0.97 \pm 0.00	60.22 \pm 0.92	0.96 \pm 0.00
		Energy-based	0.86 \pm 0.01	25.07 \pm 0.76	0.88 \pm 0.01	0.93 \pm 0.01	39.10 \pm 0.96	0.94 \pm 0.01
2251 2252 2253	RAPS	Standard	0.86 \pm 0.01	28.44 \pm 0.78	0.88 \pm 0.01	0.93 \pm 0.01	46.22 \pm 1.45	0.94 \pm 0.01
		Prevalence-Adj.	0.92 \pm 0.00	35.42 \pm 0.69	0.91 \pm 0.00	0.97 \pm 0.00	59.12 \pm 0.77	0.96 \pm 0.00
		Energy-based	0.86 \pm 0.01	25.37 \pm 0.70	0.88 \pm 0.01	0.93 \pm 0.01	39.63 \pm 0.98	0.94 \pm 0.01
2254 2255 2256	SAPS	Standard	0.86 \pm 0.01	28.26 \pm 0.76	0.88 \pm 0.01	0.93 \pm 0.01	45.73 \pm 1.28	0.94 \pm 0.01
		Prevalence-Adj.	0.92 \pm 0.00	35.69 \pm 0.66	0.91 \pm 0.00	0.97 \pm 0.00	59.11 \pm 0.95	0.96 \pm 0.00
		Energy-based	0.86 \pm 0.01	25.41 \pm 0.71	0.88 \pm 0.01	0.93 \pm 0.01	39.72 \pm 1.00	0.94 \pm 0.01
CIFAR-100-LT ($\lambda = 0.03$, severe imbalance)								
2257 2258 2259	LAC	Standard	0.84 \pm 0.01	24.43 \pm 0.63	0.87 \pm 0.01	0.92 \pm 0.01	38.72 \pm 0.77	0.94 \pm 0.01
		Prevalence-Adj. (PAS)	0.90 \pm 0.01	34.18 \pm 1.00	0.90 \pm 0.01	0.95 \pm 0.01	49.60 \pm 1.43	0.95 \pm 0.01
		Energy-based	0.84 \pm 0.01	24.06 \pm 0.64	0.87 \pm 0.01	0.92 \pm 0.01	38.04 \pm 0.90	0.94 \pm 0.01
2260 2261 2262	APS	Standard	0.85 \pm 0.01	25.64 \pm 0.94	0.88 \pm 0.01	0.92 \pm 0.01	39.43 \pm 1.23	0.94 \pm 0.01
		Prevalence-Adj.	0.95 \pm 0.00	48.47 \pm 0.65	0.93 \pm 0.00	0.98 \pm 0.00	66.59 \pm 0.73	0.97 \pm 0.01
		Energy-based	0.84 \pm 0.01	24.43 \pm 0.88	0.87 \pm 0.01	0.92 \pm 0.01	38.78 \pm 1.12	0.94 \pm 0.01
2263 2264 2265	RAPS	Standard	0.85 \pm 0.01	28.00 \pm 1.00	0.87 \pm 0.01	0.92 \pm 0.01	43.81 \pm 1.37	0.94 \pm 0.01
		Prevalence-Adj.	0.95 \pm 0.01	46.46 \pm 0.88	0.93 \pm 0.01	0.98 \pm 0.00	65.14 \pm 0.70	0.97 \pm 0.01
		Energy-based	0.84 \pm 0.01	24.74 \pm 0.69	0.87 \pm 0.01	0.92 \pm 0.01	39.74 \pm 1.17	0.94 \pm 0.01
2266 2267	SAPS	Standard	0.85 \pm 0.01	27.64 \pm 0.85	0.88 \pm 0.01	0.92 \pm 0.01	43.29 \pm 1.47	0.94 \pm 0.01
		Prevalence-Adj.	0.95 \pm 0.00	46.81 \pm 0.81	0.93 \pm 0.01	0.98 \pm 0.00	65.25 \pm 0.69	0.97 \pm 0.01
		Energy-based	0.84 \pm 0.01	24.76 \pm 0.63	0.87 \pm 0.01	0.92 \pm 0.01	39.73 \pm 1.24	0.94 \pm 0.01

2268 **L CLASS-CONDITIONAL CONFORMAL PREDICTION**
2269

2270 Beyond marginal coverage guarantee, we evaluate the proposed Energy-based nonconformity scores
2271 within the class-conditional setting. Class-conditional Conformal Prediction (CP) operates by parti-
2272 tioning the calibration dataset according to the true labels. Nonconformity score quantiles are then
2273 computed independently for each class using its respective calibration subset. The objective is to
2274 achieve *class-conditional coverage*, defined as:

2275
$$\mathbb{P}(Y_{\text{test}} \in \mathcal{C}(X_{\text{test}}) \mid Y_{\text{test}} = y) \geq 1 - \alpha, \quad \text{for all } y \in \mathcal{Y}, \quad (55)$$
2276

2277 This condition ensures that for every class $y \in \mathcal{Y}$, the probability of the true label being included in
2278 the prediction set $\mathcal{C}(X_{\text{test}})$ is at least $1 - \alpha$.

2279 We evaluate performance on the Places365 dataset at miscoverage levels $\alpha \in \{0.05, 0.1\}$ using
2280 average set size, CovGap, and SSCV (for details of these refer to E.3). As shown in Table 13, Energy-
2281 based variants consistently yield more efficient prediction sets, reflected in reduced average set sizes
2282 across all base nonconformity functions. Importantly, this improvement in efficiency does not com-
2283 promise class-conditional validity and CovGap is preserved or slightly improved in most cases.
2284 While we report SSCV for completeness—given its use in prior work—and note that it is often
2285 maintained or even improved in our experiments, we emphasize that it is not a reliable measure of
2286 conditional coverage quality.

2287 Table 13: Class-conditional performance comparison of different nonconformity score functions and their
2288 Energy-based variants on the Places365 dataset at miscoverage levels $\alpha = 0.05$ and $\alpha = 0.1$. For Set Size,
2289 CovGap, and SSCV, lower values indicate better performance. **Bold** values denote the best result within each
2290 method family. Results are averaged over 10 trials with a ResNet-50 model.

2291

Method	$\alpha = 0.1$				$\alpha = 0.05$				
	Coverage	Set Size \downarrow	CovGap \downarrow	SSCV \downarrow	Coverage	Set Size \downarrow	CovGap \downarrow	SSCV \downarrow	
APS	w/o Energy	0.89 ± 0.00	9.13 ± 0.14	5.03 ± 0.14	0.119 ± 0.015	0.95 ± 0.00	21.50 ± 0.59	3.25 ± 0.12	0.124 ± 0.039
	w/ Energy	0.89 ± 0.00	8.65 ± 0.17	5.09 ± 0.18	0.100 ± 0.00	0.95 ± 0.00	19.20 ± 0.49	3.24 ± 0.10	0.087 ± 0.027
RAPS	w/o Energy	0.89 ± 0.00	9.11 ± 0.18	4.98 ± 0.15	0.100 ± 0.00	0.95 ± 0.00	22.03 ± 0.59	3.24 ± 0.13	0.090 ± 0.024
	w/ Energy	0.89 ± 0.00	8.59 ± 0.17	4.95 ± 0.19	0.112 ± 0.011	0.95 ± 0.00	19.11 ± 0.47	3.22 ± 0.11	0.119 ± 0.020
SAPS	w/o Energy	0.89 ± 0.00	8.92 ± 0.20	4.98 ± 0.20	0.100 ± 0.00	0.95 ± 0.00	21.68 ± 0.58	3.26 ± 0.14	0.060 ± 0.019
	w/ Energy	0.89 ± 0.00	8.62 ± 0.20	4.87 ± 0.13	0.103 ± 0.01	0.95 ± 0.00	18.99 ± 0.49	3.17 ± 0.11	0.083 ± 0.013

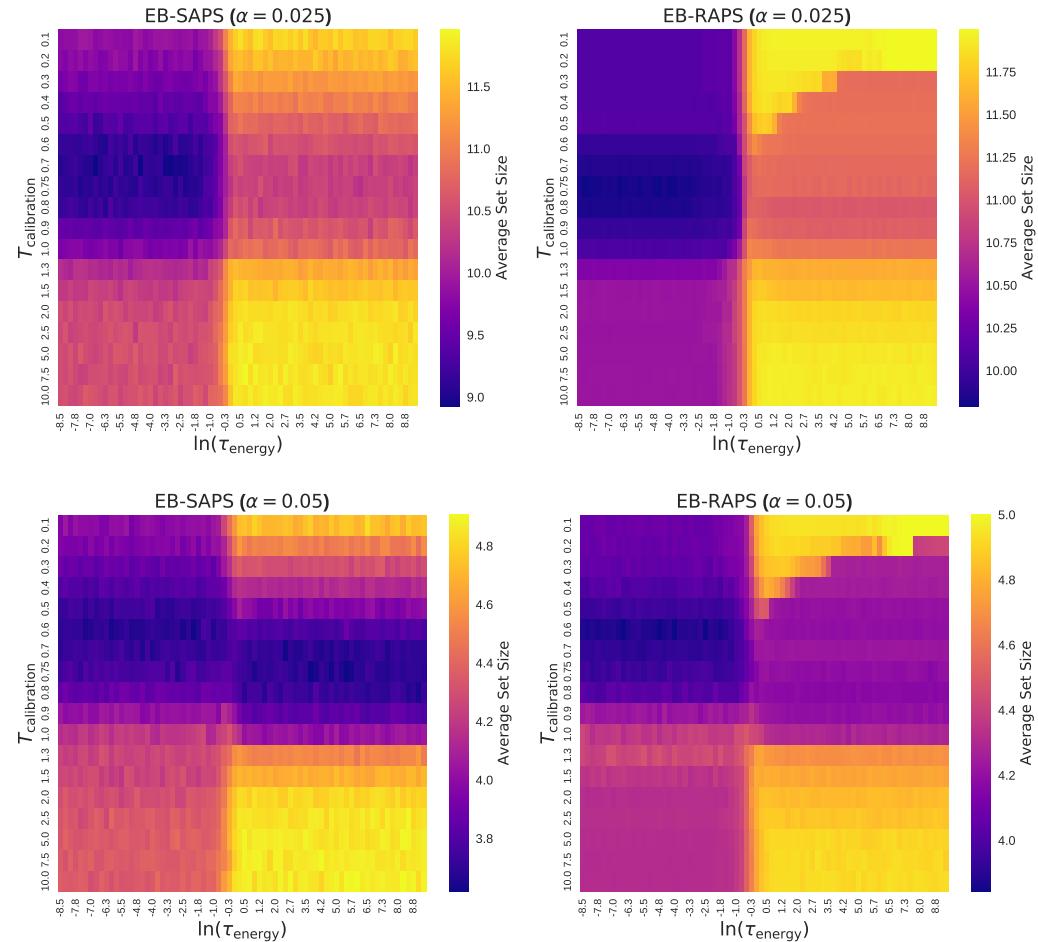
2303 **M EFFECT OF RAPS HYPERPARAMETERS λ AND k_{REG}**
2304

2305 In this section, we evaluate the sensitivity of the Regularized Adaptive Prediction Sets (RAPS)
2306 method, and its energy-based variant, to their two core hyperparameters: the regularization weight
2307 λ and the penalty threshold k_{reg} . Table 14 compares the average prediction set size of the standard
2308 RAPS baseline against our proposed Energy-based RAPS on the Places365 dataset using a ResNet-
2309 50 model, across a grid of parameter configurations.

2310 Table 14: Comparison of average prediction set sizes for varying regularization parameters (k_{reg} and λ). Lower
2311 set size is better. Values are reported as: Size (RAPS \rightarrow Energy-based RAPS). The **bold** value highlights the
2312 superior (smaller) set size.

2313

Average Set Size (w/o Energy \rightarrow w/ Energy) \downarrow							
Regularization Penalty (λ)							
k_{reg}	0	0.05	0.1	0.2	0.5	0.7	1.0
1	14.08 \rightarrow 12.88	13.11 \rightarrow 12.88	13.60 \rightarrow 12.24	14.07 \rightarrow 12.61	14.30 \rightarrow 12.90	14.30 \rightarrow 12.95	14.30 \rightarrow 13.00
2	14.09 \rightarrow 12.86	13.13 \rightarrow 12.83	13.58 \rightarrow 12.34	14.05 \rightarrow 12.63	14.30 \rightarrow 12.97	14.30 \rightarrow 13.05	14.30 \rightarrow 13.11
5	14.01 \rightarrow 12.85	13.11 \rightarrow 13.06	13.61 \rightarrow 12.67	14.07 \rightarrow 12.86	14.30 \rightarrow 13.22	14.30 \rightarrow 13.29	14.30 \rightarrow 13.34
10	14.05 \rightarrow 12.90	13.13 \rightarrow 13.07	13.58 \rightarrow 13.57	14.07 \rightarrow 13.46	14.29 \rightarrow 13.62	14.30 \rightarrow 13.74	14.30 \rightarrow 13.82
50	14.00 \rightarrow 12.86	13.97 \rightarrow 12.91	14.02 \rightarrow 12.85	14.00 \rightarrow 12.85	14.02 \rightarrow 12.88	14.05 \rightarrow 12.87	13.90 \rightarrow 12.88

2322 N EFFECT OF $T_{\text{CALIBRATION}}$ AND τ_{ENERGY}
23232355
2356 Figure 8: Average Set Size heatmap for different hyperparameter settings across Energy-based variants of
2357 RAPS and SAPS.

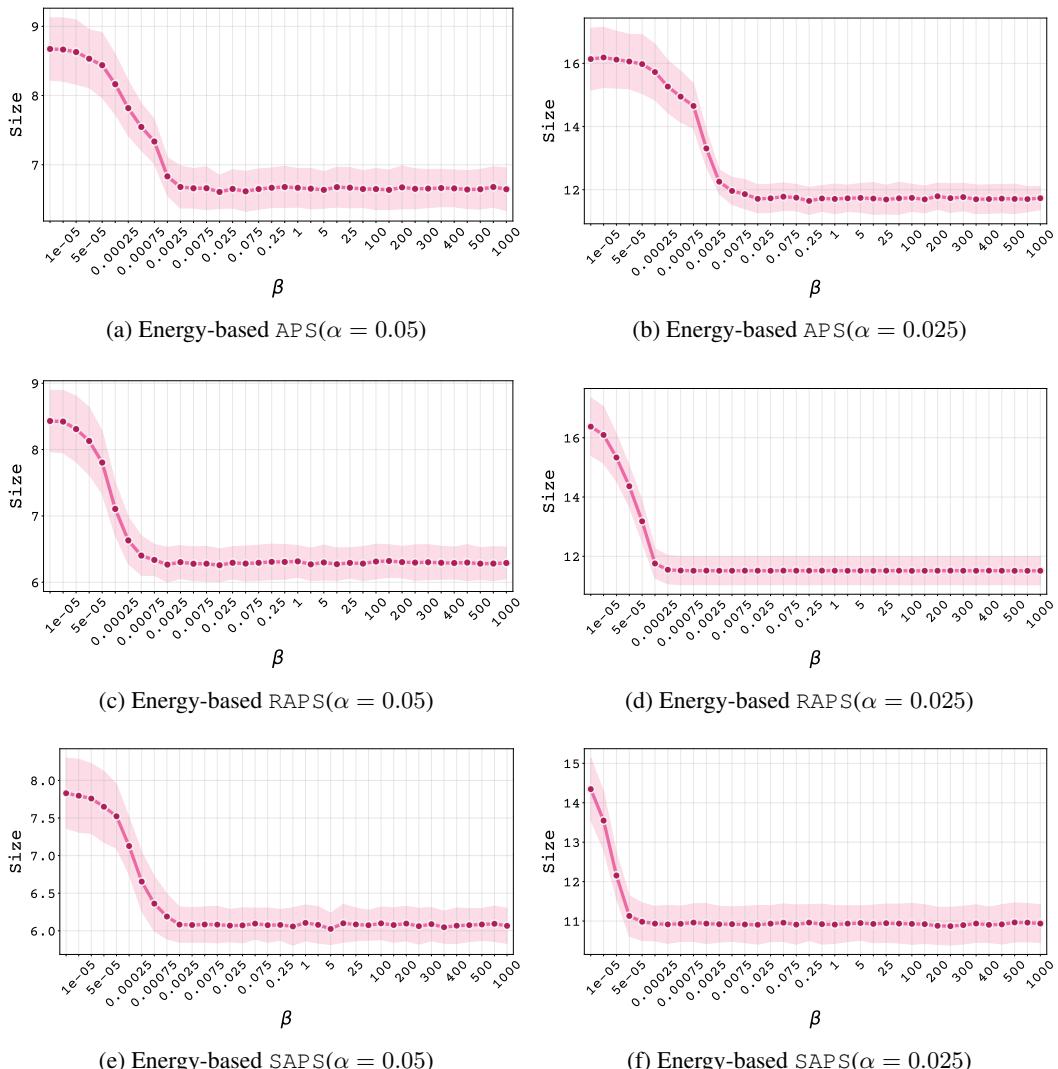
2358 Sensitivity heatmaps illustrate how T (temperature in the calibrated softmax) and τ (temperature in the energy calculation) affect the average set size on the ImageNet dataset for Energy-Based
2359 RAPS and Energy-Based SAPS, with $\alpha \in \{0.025, 0.05\}$. As $\ln(\tau)$ increases, the effect of energy-
2360 based reweighting gradually diminishes. Consequently, for larger values of τ , the model converges
2361 to the baseline method. For instance, Energy-Based RAPS with a large positive $\ln(\tau)$ behaves
2362 almost identically to standard RAPS. As shown in Figure 8, across different values of T (softmax
2363 probability calibration), energy-based variants of the methods (corresponding to smaller values of
2364 $\ln(\tau)$) produce more informative prediction sets compared to their baseline counterparts (associated
2365 with larger values of $\ln(\tau)$).

2376 O SENSITIVITY ANALYSIS OF THE SOFTPLUS PARAMETER β 2377

2378 We analyze the impact of the sharpness parameter β on the efficiency of the generated prediction
 2379 sets. As defined in Equation 9, this parameter controls the approximation of the scaling factor to the
 2380 ReLU function. We evaluate the average prediction set size across a wide range of β values for the
 2381 CIFAR-100 dataset using a ResNet-56 model.

2382 Figure 9 presents the results for Energy-based APS, Energy-based RAPS, and Energy-based SAPS at
 2383 miscoverage levels $\alpha = 0.05$ and $\alpha = 0.025$. As β approaches zero, the term $\frac{1}{\beta} \log(1 + e^{-\beta F(x)})$
 2384 yields scaling factors that are inseparable across samples. Due to this loss of distinction, the per-
 2385 formance converges to the baseline without energy.

2386 However, as β increases, the performance stabilizes and remains constant across several orders of
 2387 magnitude. This behavior aligns with the theoretical motivation that the scaling factor need only ap-
 2388 proximate the ReLU function to handle rare negative free energy values while preserving the signal
 2389 for positive values. Consequently, precise tuning of this parameter is unnecessary. Selecting a suf-
 2390 ficiently large value is a safe option to achieve the performance benefits of Energy-based conformal
 2391 classification.



2427 Figure 9: Ablation study of the parameter β on CIFAR-100 with ResNet-56. The plots show the average
 2428 prediction set size (shaded regions indicate standard deviation) as a function of β . Performance stabilizes for
 2429 sufficiently large values of β , indicating that the method does not require specific tuning of this parameter.

2430 **P ON THE RELIABILITY OF CONFORMAL CLASSIFIERS WHEN FACED WITH**
 2431 **OOD TEST SAMPLES**
 2432

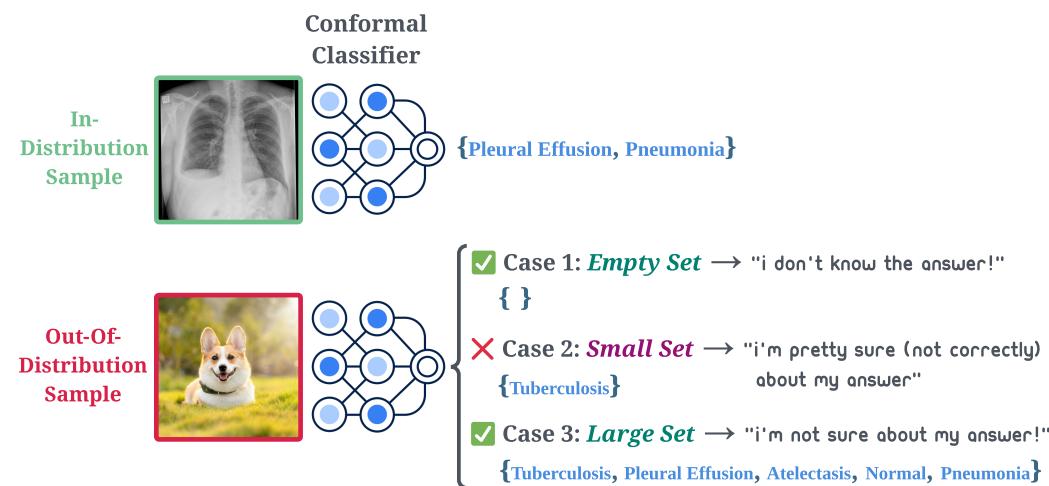
2433 A reliable system for uncertainty quantification is often expected to satisfy two primary objectives,
 2434 as outlined in the Appendix of [Angelopoulos & Bates \(2021\)](#):

2435

- 2436 1. Flag out-of-distribution (OOD) inputs to avoid making predictions on unfamiliar data.
- 2437 2. If an input is deemed in-distribution, output a prediction set that contains the true class with
 2438 user-specified probability.

2439
 2440 A practical strategy to achieve this is a two-stage pipeline: first, an OOD detector screens each input.
 2441 If an input is flagged as OOD, the system can abstain (e.g., by returning an empty set). If deemed
 2442 in-distribution, the input is passed to a conformal predictor to generate a valid prediction set. This
 2443 separation, while effective, requires deploying and managing two distinct models.

2444 The utility of the energy-based paradigm becomes particularly evident in scenarios where a dedicated
 2445 OOD detection module is not available to filter inputs. In such cases, a standard conformal
 2446 predictor, relying solely on softmax outputs, can be misleading. An OOD input might still produce
 2447 a single, high-confidence softmax score, leading the base conformal method to output a small, high-
 2448 confidence prediction set (e.g., {‘Tuberculosis’}) for a non-medical image. This false confidence is
 2449 a critical failure mode.



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 Figure 10: Conceptual diagram of a reliable conformal classifier facing an OOD input. The desired behaviors are to produce an empty or large set, both signaling uncertainty, and to avoid producing a small set that implies false confidence.

The energy-based approach addresses this vulnerability by incorporating a more reliable measure of model uncertainty. The key difference in behavior is:

- For *familiar ID inputs*, model confidence is high, resulting in a high negative free energy. The prediction sets are thus appropriately small and efficient, similar to the base variant.
- For *unfamiliar OOD inputs*, the model’s uncertainty is captured by a low negative free energy. The energy-based reweighting dampens scores so that more classes fall below the calibrated quantile. This results in the predictor generating a large prediction set.

This is a clear improvement over the base variant, which is prone to producing small, overconfident sets for such inputs. Therefore, even when the system is not configured to abstain, the large prediction set generated by the energy-based method provides a more robust and honest signal of uncertainty. It reduces the risk of overconfident and incorrect predictions on OOD data, making it a more reliable choice in deployments without a separate OOD detector. This property validates

2484 the use of free energy as a model-aware signal that overcomes the limitations of standard softmax
 2485 probabilities, as noted in prior work (Liu et al., 2020; Wang et al., 2021).
 2486

2487 To illustrate the importance of the desiderata outlined in Section 3.3, consider a conformal classifier
 2488 trained to identify medical conditions from chest X-rays. If this model is fed with an image of a
 2489 completely unrelated subject, such as a corgi, a reliable classifier must signal its unfamiliarity with
 2490 the input. As visualized in Figure 10, this signal correctly manifests in two ways:

- 2491 1. An **empty set** (\emptyset), which communicates: “*I don’t know the answer.*”
- 2492 2. A **large set** (e.g., {‘Tuberculosis’, ‘Pleural Effusion’, ‘Atelectasis’, ...}), which communi-
 2493 cates: “*I am uncertain about the answer.*”

2494 In contrast, producing a small, non-empty set (e.g., {‘Tuberculosis’}) is misleading, as it incorrectly
 2495 signals high confidence in a prediction that is likely wrong.

2496 Our experimental results in Section 3.3 confirm this behavior in practice. We observed an increase
 2497 in the average prediction set size for the energy-based variants of the nonconformity scores.
 2498

2500 Q ENERGY-BASED REWEIGHTING VS. ENTROPY-BASED REWEIGHTING

2501 A common measure of uncertainty in a classifier’s output is the *Shannon Entropy* of its softmax
 2502 probability distribution. As described by Luo & Colombo (2024), let $\mathbf{f}(x)$ be the logit vector pro-
 2503 duced by the classifier for an input x , and let $\hat{\pi}(k | x)$ be the softmax probability for class k . The
 2504 Entropy $\mathbb{H}(x)$ is given by:

$$\begin{aligned}
 2507 \mathbb{H}(x) &= - \sum_{k=1}^K \hat{\pi}(k | x) \log \hat{\pi}(k | x) \\
 2508 &= - \sum_{k=1}^K \hat{\pi}(k | x) \left(f_k(x) - \log \sum_{j=1}^K \exp(f_j(x)) \right) \\
 2509 &= - \sum_{k=1}^K \hat{\pi}(k | x) f_k(x) + \left(\log \sum_{j=1}^K \exp(f_j(x)) \right) \sum_{k=1}^K \hat{\pi}(k | x) \\
 2510 &= - \sum_{k=1}^K \hat{\pi}(k | x) f_k(x) + \log \sum_{j=1}^K \exp(f_j(x)) \tag{56}
 \end{aligned}$$

2511 If one were to consider an “Entropy-based reweighting” for conformal scores, it would likely utilize
 2512 this $\mathbb{H}(x)$ or a function thereof. However, the decomposition in Equation 56 reveals two distinct
 2513 components influencing the Entropy value. The first term, $-\sum_{k=1}^K \hat{\pi}(k | x) f_k(x)$, depends on the
 2514 alignment of softmax probabilities with the logit values. The second term is the `logsumexp` (LSE)
 2515 of the logits: $L(x) = \log \sum_{j=1}^K \exp(f_j(x))$.
 2516

2517 This LSE term, $L(x)$, is particularly relevant as it is directly related to the concept of free energy,
 2518 which forms the basis of our proposed Energy-based nonconformity scores. It captures the overall
 2519 magnitude or scale of the raw logits. Critically, while the softmax probabilities $\hat{\pi}(k | x)$ also depend
 2520 on these logits, the LSE term $L(x)$ is calculated purely from the logits $f_j(x)$ without $\hat{\pi}(k | x)$
 2521 appearing as explicit factors within its own sum, unlike in the definition of entropy.

2522 Indeed, softmax entropy is not well-suited for capturing epistemic uncertainty (Mukhoti et al., 2021).
 2523 The distinction and potential advantage of using an Energy-based measure over the Entropy $\mathbb{H}(x)$
 2524 is illustrated in Figure 2. As shown, both the “easy” and “hard” samples can yield high softmax
 2525 confidence for the predicted class, resulting in very low Entropy values (close to zero) for both.
 2526 This suggests that Entropy alone might not adequately distinguish between an input for which the
 2527 model is genuinely certain (high overall logit values, “easy sample”) and an input where the model
 2528 is less certain overall but still produces a peaky softmax distribution (“hard sample” with high softmax
 2529 for one class). In contrast, the negative energy scores clearly differentiate these two cases:

the easy sample exhibits a significantly higher negative energy score (32.49) compared to the hard sample (14.08). This indicates that the Energy-based metric, by reflecting the overall scale of logit activation, provides a more nuanced signal of the model's underlying certainty.

Our proposed Energy-based Nonconformity Scores in Section 2.3 leverages this energy signal by reweighting a standard nonconformity score with $\text{softplus}(-F(x))$. The rationale is that $F(x)$ offers a more direct and potentially more robust indication of the model’s overall certainty about an input x than the Entropy $\mathbb{H}(x)$, which can be saturated (i.e., near zero) for different levels of underlying model certainty. Empirical comparisons supporting the benefits of Energy-reweighted scores over Entropy-reweighted alternatives is provided in Table 15.

Table 15: Performance comparison of different nonconformity score functions and their Energy-based and Entropy-based variants on ImageNet using a ResNet-50 classifier at miscoverage levels $\alpha = 0.05$ and $\alpha = 0.1$. Results are averaged over 10 trials and reported as empirical coverage and average prediction set size. **Bold** values indicate the best performance within each group. The Entropy-based variants of adaptive method, APS, RAPS, and SAPS are defined as $S_{\mathbb{H}}(x, y) = \frac{S(x, y)}{\mathbb{H}(x)}$, and for LAC it is defined as $S_{\mathbb{H}}(x, y) = -\hat{\pi}(y|x) \cdot \mathbb{H}(x)$.

		$\alpha = 0.1$		$\alpha = 0.05$	
Family		Coverage	Set Size ↓	Coverage	Set Size ↓
LAC	baseline	0.898 ± 0.002	1.487 ± 0.013	0.950 ± 0.002	2.682 ± 0.039
	w/ Energy	0.898 ± 0.002	1.485 ± 0.012	0.949 ± 0.002	2.680 ± 0.043
	w/ Entropy	0.898 ± 0.002	1.496 ± 0.014	0.949 ± 0.002	2.696 ± 0.043

APS	baseline	0.899 ± 0.002	1.605 ± 0.022	0.950 ± 0.002	4.007 ± 0.164
	w/ Energy	0.899 ± 0.002	1.599 ± 0.022	0.950 ± 0.002	3.842 ± 0.159
	w/ Entropy	0.898 ± 0.003	2.159 ± 0.042	0.949 ± 0.002	4.990 ± 0.083

RAPS	baseline	0.898 ± 0.003	1.764 ± 0.030	0.949 ± 0.001	4.222 ± 0.056
	w/ Energy	0.898 ± 0.003	1.763 ± 0.033	0.949 ± 0.001	3.889 ± 0.057
	w/ Entropy	0.898 ± 0.003	2.108 ± 0.052	0.949 ± 0.002	4.811 ± 0.076

SAPS	baseline	0.898 ± 0.002	1.664 ± 0.034	0.949 ± 0.002	3.659 ± 0.073
	w/ Energy	0.898 ± 0.002	1.662 ± 0.029	0.949 ± 0.002	3.654 ± 0.064
	w/ Entropy	0.898 ± 0.003	2.101 ± 0.039	0.949 ± 0.002	4.702 ± 0.075

2592 **R DETAILED RESULTS ON IMAGENET**
25932594 The table 16 provides an evaluation of 3 nonconformity score across 16 different model architectures
2595 on the ImageNet validation set. We compare adaptive baseline methods (APS, RAPS, and SAPS)
2596 against their respective Energy-based counterparts.
25972598 Table 16: Comparison of average prediction set sizes and accuracy for conformal methods on ImageNet at
2599 confidence level of 95% ($\alpha = 0.05$). Lower average set size is better. Set sizes are shown as baseline (w/o
2600 Energy) \rightarrow Energy-based (w/ Energy). The **bold** value highlights the superior (smaller) set size within each
2601 pair. The results are reported as the median of means over 10 trials.
2602

Model	Accuracy		Average Set Size (w/o Energy \rightarrow w/ Energy) \downarrow		
	Top-1	Top-5	APS	RAPS	SAPS
ResNet152	0.783	0.940	3.82 \rightarrow 3.48	3.25 \rightarrow 3.17	2.87 \rightarrow 2.83
ResNet101	0.774	0.936	3.97 \rightarrow 3.70	3.60 \rightarrow 3.32	3.15 \rightarrow 3.03
ResNet50	0.761	0.929	4.09 \rightarrow 3.97	4.16 \rightarrow 3.84	3.70 \rightarrow 3.63
ResNet34	0.733	0.914	9.86 \rightarrow 8.29	9.94 \rightarrow 7.73	9.57 \rightarrow 7.41
ResNet18	0.698	0.891	14.23 \rightarrow 12.06	15.16 \rightarrow 11.43	14.56 \rightarrow 10.93
VGG19	0.742	0.918	8.41 \rightarrow 7.14	8.85 \rightarrow 6.85	8.36 \rightarrow 6.71
VGG16	0.734	0.915	8.70 \rightarrow 7.34	9.06 \rightarrow 7.01	8.67 \rightarrow 6.89
VGG13	0.716	0.904	10.77 \rightarrow 9.29	11.83 \rightarrow 8.78	11.60 \rightarrow 8.58
VGG11	0.704	0.898	12.36 \rightarrow 10.55	13.42 \rightarrow 10.09	13.10 \rightarrow 9.62
ViT-B/16	0.811	0.953	4.70 \rightarrow 4.29	4.10 \rightarrow 3.56	3.53 \rightarrow 3.24
ViT-B/32	0.759	0.925	9.62 \rightarrow 8.25	8.52 \rightarrow 7.28	7.93 \rightarrow 6.68
Swin_s	0.832	0.964	2.79 \rightarrow 2.81	3.13 \rightarrow 2.87	2.74 \rightarrow 2.65
Swin_t	0.815	0.958	3.36 \rightarrow 3.38	3.64 \rightarrow 3.42	3.28 \rightarrow 3.18
EfficientNet_b4	0.834	0.966	5.87 \rightarrow 4.99	4.87 \rightarrow 4.28	4.33 \rightarrow 3.93
EfficientNet_v2_m	0.851	0.972	5.93 \rightarrow 5.60	5.16 \rightarrow 5.16	4.86 \rightarrow 4.69
ShuffleNet_v2_x1_0	0.694	0.883	19.17 \rightarrow 14.79	19.63 \rightarrow 14.48	19.40 \rightarrow 14.07
Average	0.765	0.929	7.98 \rightarrow 6.87	8.02 \rightarrow 6.45	7.60 \rightarrow 6.13

2628
2629 **S AI USAGE CLARIFICATION**
26302631 Large Language Models were used to polish the writing of this manuscript by improving grammar,
2632 spelling, sentence flow, and overall readability. All research design, analysis, and interpretation were
2633 fully carried out and decided by the authors.
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