Relation-Oriented: Toward Causal Knowledge-Aligned AGI

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Abstract

The current relationship modeling paradigm, grounded in the observational i.i.d assumption, fundamentally misaligns with our causal knowledge understanding due to two key oversights: 1) the unobservable relations, which lead to undetectable hierarchical levels of knowledge, driving the need for model generalizability; 2) the cognitive relative timings, which crucially support our structural knowledge comprehension, resulting in inherent biases within the present *Observation-Oriented* paradigm. Adopting a novel *Relation-Oriented* perspective, this paper proposes a new framework to unify the various confusions surrounding causality learning, ranging from traditional causal inference to modern language models.

Also, relation-indexed representation learning (RIRL) is raised as a baseline implementation method of the proposed new paradigm, alongside comprehensive experiments demonstrating its efficacy in autonomously identifying dynamical effects in relationship learning.

1 Introduction

The concept of Artificial General Intelligence (AGI) has prompted extensive discussions over the years Newell (2007), with the central target toward facilitating human-like causal reasoning and comprehension in AI systems Marcus (2020). In recent years, the large language models (LLMs) have risen as notable achievements in language-understanding tasks and accordingly evoked debates about whether LLMs have edged us closer to realizing AGI Rylan (2023). Some studies point to their shortcomings in truly comprehending causality Pavlick (2023), while others argue in favor of LLMs' ability to represent complex spatial and temporal features Wes (2023). Notably, the use of meta-learning in language models has shown potential in achieving human-like generalization capabilities, at least to a certain extent Lake (2023).

These debates are anchored in an underlying inquiry: What underpins the distinction between two types of generalization? One is how humans generalize learned causal knowledge to diverse scenarios, and another is how AI systems generalize associative knowledge among texts and images.

It appears that classical causal inference has offered a clear delineation between causality and mere correlations or associations Pearl et al. (2000); Peters et al. (2017). Moreover, it has established a robust theoretical groundwork for representing causality in computational models - Building on which, widely utilized causal learning techniques have yielded significant contributions to the accumulation of causal knowledge in various fields Wood (2015); Vuković (2022); Ombadi (2020). It is thus logical to incorporate existing causal knowledge, often represented as causal DAGs (Directed Acyclic Graphs), into AI model architectures Marwala (2015); Lachapelle et al. (2019). While this integration has greatly enhanced learning efficiency, it has not yet achieved the level of generalizability that constitutes a success Luo (2020); Ma (2018).

This likely circles us back to the initial question, as causal inference cannot directly bridge the gap between human-like causal reasoning and current AI systems. However, it does offer a different perspective: How would humans conduct causal reasoning based solely on DAGs? A task that evidently challenges AI.

Even within the realm of causal inference, the process of converting DAGs into operational causal models is rigorous Elwert (2013). Tailored adjustments and interpretations are often required, reliant on human discernment across varied applications Sanchez (2022); Crown (2019). Key challenges include establishing fundamental causal assumptions Sobel (1996), addressing confounding effects Greenland (1999), ensuring model interpretability Pearl et al. (2000), etc. These achievements constitute the cornerstone of the value provided by causal inference methodologies. It stands to reason that the answer to this fundamental question may be gleaned from examining the challenges that causal inference has faced and partially overcome.

From an applicational standpoint, Scholkopf (2021) have synthesized the development of current causal models, underscoring the pivotal role of realizing "causal representations" to achieve the generalizability of AI-based causal models across different "levels of knowledge" learning. They propose the potential need for a "new learning paradigm" - an idea we find both logical and thought-provoking. Our current models, ranging from causal to AI, are chiefly based on the assumption of independent and identically distributed (i.i.d.) observations, a paradigm that may be hindering their ability to autonomously realize generalizable causal reasoning. On the other hand, Zhang (2012) points out the "identifiability difficulty" when facing nonlinear effects, an inherent obstacle under the i.i.d observational effect setting.

For clarity, we designate the prevailing paradigm as ***Observation-Oriented*** modeling. In this study, we propose a novel paradigm, termed ***Relation-Oriented*** modeling, inspired by the relation-indexing nature of human cognition processes Pitt (2022). Through this new lens, we seek to pinpoint the intrinsic limitations of the existing paradigm. Accordingly, to validate the proposed new paradigm, it must shed light on the array of questions that have emerged from the outset. To encapsulate these queries:

- *Firstly*, causal inference challenges such as confounding effects, dependency on causal assumptions, and interpretative complexities call for a foundational explanation.
- *Secondly*, To integrate causal reasoning within AI models, we need a nuanced understanding of "levels of knowledge," the essential role of causal representation, its relevance to the difficulty of identifying nonlinear effects, and potential resolutions to these issues.
- *Thirdly*, in the context of Large Language Models (LLMs), we must discern the distinction between the "spatial and temporal" conceptions in language versus causality comprehension, and critically assess what meta-learning has accomplished in terms of generalizability.

While these questions might seem disparate, they are intrinsically linked to the fundamental requirement by the *Observation-Oriented* paradigm: it necessitates the prior specification of observable entities (including temporal events). In solely observational learning tasks (like image recognition), these entities serve as the modeling target. In causal relationship learning, they are priorly identified as causes and effects, with their interrelation acting as the primary learning objective.

This fundamental requirement introduces two ***primary limitations*** in modeling: 1) the inability to account for unobservable relational knowledge, which leads to undetectable hierarchical levels in modeling, and 2) the obligation to assign timestamps to events, potentially causing the overlook of relative timings underpinning structuralized dynamics in our relational knowledge, and essentially introducing inherent biases.

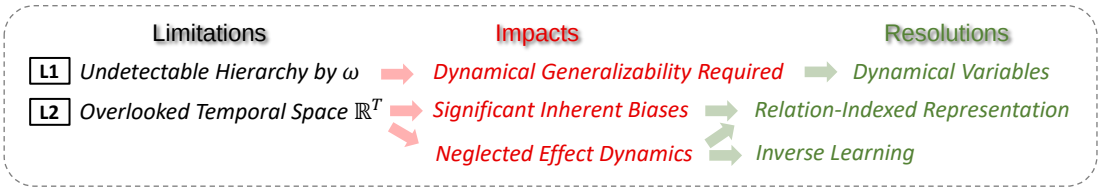


Figure 1: Overview of the *Observation-Oriented* paradigm's primary limitations (labeled as L1 and L2), featuring the concept of hidden relation ω (see section 1.2) and the temporal space $\mathbb{R}^T$ with relative timing axes (discussed in section 2.1).

This paper consists of four principal parts:

1. the Introduction, which sets the foundation for the proposed *Relation-Oriented* perspective in section 1.1, and analyze the roles of unobservable relational knowledge in modeling, using an illustrative example to explain its resulting undetectable hierarchy in section 1.2 (i.e., the limitation L1).
2. Chapter I, including Sections 2 through 4, establishes the *Relation-Oriented* framework to decompose relationship modeling from a more precise perspective, and through this framework, examines the fundamental impacts of the outlined limitations, and addresses the queries listed above.
3. Chapter II, from Sections 5 to 7, introduces *Relation-Indexed Representation Learning* (RIRL) as a baseline realization method of the *Relation-Oriented* paradigm and evaluates the efficacy of relation-indexed autonomous effect identification.
4. the Conclusion in Section 8 summarizes the insights and findings of this study.

1.1 Relation-Oriented Perspective

Typically, experiments with n trials produce instances $x^n = x_1, \ldots, x_n$ from sequential random variables $X^n = X_1, \ldots, X_n$, which are usually assumed to be independent and identically distributed (i.i.d.). When these variables evolve over time, n is often replaced by the timestamp t to get temporal sequence $X^t = X_1, \ldots, X_t$, maintaining the i.i.d. assumption, and the relationship function is in shape $Y = f(X^t; \theta)$.

In our research, we abandon the i.i.d. assumption over $\{X_i \mid i = 1, \ldots, t\}$ in the temporal dimension $\mathbf{t}$, instead treat their sequence X^t as a single entity, denoted by variable $\mathcal{X} \in \mathbb{R}^{d+1}$, with d representing the observational dimension of each instance X_i. For clarity, we use $X \in \mathbb{R}^d$ to represent a solely observational variable, and let $\mathcal{X} = \langle X, \mathbf{t} \rangle \in \mathbb{R}^{d+1}$ derived by incorporating the $\mathbf{t}$-dimension to encompass features across both observational and temporal dimensions. It is worth noting that variables such as $\mathcal{X}$ are conventionally referred to as spatial-temporal Andrienko (2003). However, in this context, "spatial" is broadly interpreted to mean "observational" and is not restricted to physical spatial data, such as geographic coordinates.

Consider the functional relationship $\mathcal{Y} = f(\mathcal{X}; \theta)$, where $\mathcal{Y} = \langle Y, \tau \rangle \in \mathbb{R}^{b+1}$ with τ representing the temporal evolution of $Y \in \mathbb{R}^b$. We employ the Fisher Information $\mathcal{I}_{\mathcal{X}}(\theta)$ Ly et al. (2017) of $\mathcal{X}$ about θ, to define the component of $\mathcal{Y}$ (signified as $\hat{\mathcal{Y}}$) that is sufficiently identified by indexing through θ:

Definition 1. the *Relation-Indexed Representation* $\hat{\mathcal{Y}}_\theta$ in Relationship Modeling.

Let the ***relation*** θ adequately represents the influence of $\mathcal{X}$ on $\mathcal{Y}$, denoted as $\mathcal{X} \xrightarrow{\theta} \mathcal{Y}$, then $\hat{\mathcal{Y}}_\theta = f(\mathcal{X}; \theta)$ represents the *sufficient* component of $\mathcal{Y}$ about θ, which is, $\mathcal{I}_{\hat{\mathcal{Y}}_\theta}(\theta) = \max\ \mathcal{I}_{\hat{\mathcal{Y}}}(\theta) = \mathcal{I}_{\mathcal{X}}(\theta)$.

Consequently, $\hat{\mathcal{Y}}_\theta$ encapsulates the information within $\mathcal{Y}$ that is entirely derived from $\mathcal{X}$, thus defined as the *relation-indexed representation*. Accordingly, the remaining component of $\mathcal{Y}$, expressed as $\mathcal{Y} - \hat{\mathcal{Y}}_\theta$, does not depend on θ. The *Relation-Oriented* perspective focuses on building models by concentrating on θ.

The notation "$\rightarrow$" typically denotes causality, although a directional relationship does not always imply causality in logic. Nonetheless, for clarity, we will adopt terminology consistent with causal inference: for relationship $\mathcal{X} \xrightarrow{\theta} \mathcal{Y}$, we refer to $\mathcal{X}$ as the *cause* and $\mathcal{Y}$ as the *effect*, with a *relation* θ connecting them. Accordingly, the definition of $\hat{\mathcal{Y}}_\theta$ is aligned with the "causal representation" concept Scholkopf (2021). Crucially, in this research, both *causality* and *correlation* denote types of relationships with a relation θ (their difference will be discussed later), while *association* (typically nonlinear) refers to statistical dependency between entities without an informative θ, expressed as $(\mathcal{X}, \mathcal{Y})$.

Remark 1. Given $\mathcal{X} \xrightarrow{\theta} \mathcal{Y}$ with ***observables*** $\mathcal{X}$ and $\mathcal{Y}$, the relationship model $\mathcal{Y} = f(\mathcal{X}; \theta)$ becomes *informative* due to the ***unobservable*** θ.

The principle outlined in Remark 1 has its origins in the concept of Common Cause Dawid (1979); Scholkopf (2021), suggesting that any nontrivial (i.e., informative) conditional independence between two observables requires a third, mutual cause (i.e., the unobservable "relation" in our context).

$\mathcal{X}$ and $\mathcal{Y}$ can be either solely observational entities, equal to X and Y (e.g., images, spatial coordinates of a quadrotor, etc.), or observational-temporal entities (e.g., trends of stocks, a quadrotor's trajectory, etc.). Regardless of their characterization, the primary goal of utilizing the function $\mathcal{Y} = f(\mathcal{X}; \theta)$ is to encapsulate the unobservable relational knowledge represented by θ, rather than merely associative distribution $(\mathcal{X}, \mathcal{Y})$.

To clarify the concept of informative θ, let's consider a simple example. In the relationship "Bob (represented as X) has a son named Jim (represented as Y)", the father-son relation information $\mathcal{I}(\theta)$ between them is evident to human cognition but unobservable to AI provided sufficiently observed social activities. Also, θ can be seen as the common cause of X and Y that makes their connection unique, rather than any random pairing of "Bob" and "Jim". Through the observational data, AI might deduce a particular associative pattern over (X, Y), but cannot internalize the unobservable information $\mathcal{I}(\theta)$ between them.

Drawing on the symbolization provided in Definition 1, a comprehensive *Relation-Oriented* framework is introduced in Section 2, offering more complete insights into the modeling of causal relationships.

1.2 Unobservable Relational Knowledge

Unobservable knowledge may not directly serve as the learning objective relation θ, but it can still be relative to and profoundly impact the modeling process. We elucidate this with the following example: It is notable that on social media, AI-created personas can have realistic faces but seldom showcase hands. This is because AI for visual tasks struggles with the intricate structure of hands, instead treating them as arbitrary assortments of finger-like items. Figure 2(a) provides AI-created hands with faithful color but unrealistic shapes, while humans can effortlessly discern hand gestures from the grayscale sketches in (b).

Humans intuitively employ informative relations as the ***indices***, guiding us to specific mental representations Pitt (2022). As illustrated in Figure 2(b), our cognition operates hierarchically, progressing through a series of relations, denoted as $\theta = \{\theta_i, \theta_{ii}, \theta_{iii}\}$. Each higher-level understanding builds upon conclusions drawn at preceding levels. Specifically, Level **I** identifies individual fingers; Level **II** distinguishes gestures based on the positions of the identified fingers, incorporating additional information from our understanding of how fingers are arranged to constitute a hand, denoted by ω_i; and Level **III** grasps the meanings of these gestures from memory, given additional information ω_{ii} from knowledge.

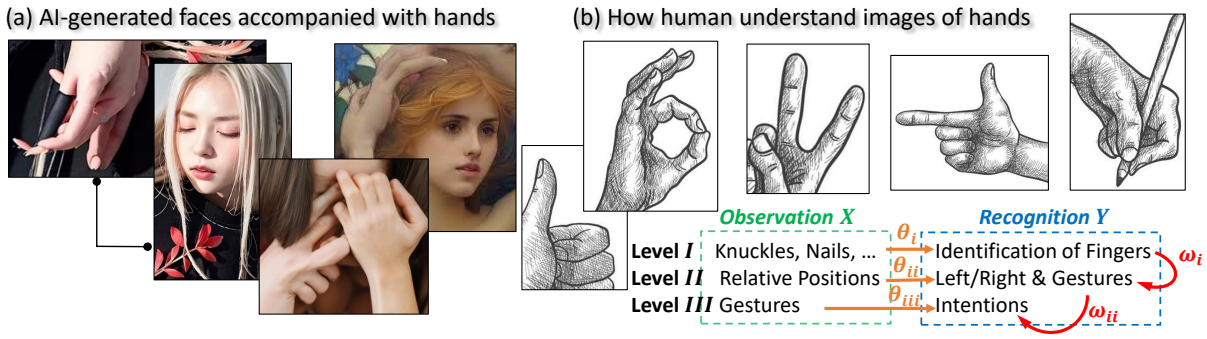


Figure 2: Unobservable relations $\theta = \{\theta_i, \theta_{ii}, \theta_{iii}\}$ and $\omega = \{\omega_i, \omega_{ii}\}$. AI can generate reasonable faces but treat hands as arbitrary mixtures of fingers; while human cognition processes observations hierarchically to avoid this mess, by indexing through a series of relations $\{\theta_i, \theta_{ii}, \theta_{iii}\}$.

Typically, these visual learning tasks do not aim to model relations, neither θ nor ω. Instead, they focus on capturing observational entities (pertaining solely to X). Without relation-indexing through θ, AI systems may not distinguish entities across different levels but only capture their associative dependence, like $(X_{II} \mid X_I)$ and $(X_{III} \mid X_I, X_{II})$, without deeper, informative insights into ω.

However, for such solely observational learning, the hidden ω may not always be essential. If entities across levels are observationally distinct and non-overlapping, AI can accurately differentiate them. For instance, AI can generate convincing faces because the appearance of eyes strongly indicates facial angle, removing the need to distinguish "eyes" $= X_{II}$ from "faces" $= X_I$. Additionally, based on fully-captured levels, AI can inversely uncover the hidden ω using methods such as reinforcement learning Sutton (2018); Arora (2021) - In this case, approvals of generated five-fingered hands may lead AI to identify fingers autonomously.

Definition 2. *Hidden Relation* ω and its resulting *Undetectable Hierarchy*.
Different from the ***indexing*** *relation* θ, the ***hidden*** *relation* ω can constitute *undetectable hierarchical levels* of knowledge, requiring model *generalizable* to be effective across.

A *generalizable* model enables the learned lower-level relationships to be reusable for higher-level learning tasks Scholkopf (2021), which mirrors our inherent capability to generalize knowledge in cognition. For example, our ability to identify fingers can be applied regardless of the types of medium, like images, photos, or videos. Conversely, generalizability also denotes the capacity to *individualize* from higher to lower levels, accommodating different ω values.

The illustration in Figure 2 highlights the distinct roles of unobservable relations (θ and ω) in modeling. Our central concern, however, is modeling relationships with θ as the primary ***objective*** *relation* for learning. In this context, ω stratifies unobservable θ into hierarchical levels, culminating in a completely imperceptible joint distribution of (θ, ω), which precludes methods like inverse reinforcement learning.

For instance, consider family incomes X influence grocery shopping frequencies Y through relation θ. Here, the cultural background ω emerges as an important factor, such that an effective model $Y = f(X; \theta)$ has to be individualizable, i.e., conditioned on a specific country (represented by a particular ω value) to ensure practical utility. On the opposite, a generalization would imply $\omega = \varnothing$.

For the sake of clarity, hereafter in this paper, unless explicitly stated otherwise, the hidden relation ω represents two hierarchical modeling levels: the generalized level $X_o \xrightarrow{\theta_o} Y_o$ with θ_o implying $\omega = \varnothing$, and the individualized level $X_\omega \xrightarrow{\theta_\omega} Y_\omega$ given θ_o with a specific ω value, collectively notated as $(\theta, \omega) = \begin{pmatrix} \theta_o \\ \theta_\omega \end{pmatrix}$.

Chapter I: Limitations of Current Observation-Oriented Paradigm

The prevalent *Observation-Oriented* modeling paradigm inherently misaligns with the relation-centric human comprehension Pitt (2022). This misalignment may not have been critical in the past. In traditional causal inference, challenges could be addressed through intended adjustments due to the limited scale of questions. Nonetheless, with the advancements in AI-based large models, the consequences of this misalignment have become increasingly significant across various applications.

Section 2 establishes a *Relation-Oriented* dimensionality framework to symbolize causal relationship models; through which, we recognize the critical role of relative timings (highlighted as limitation $\boxed{\text{L2}}$), and explore the essence of generalizability for a relationship model comprising structuralized dynamics. Subsequently, Section 3 delves into the critical implications of overlooked effect dynamics (the secondary impact of $\boxed{\text{L2}}$), and accordingly reevaluates the traditional causal inference challenges within the new framework. Lastly, Section 4 elucidates the inherent biases that *Observation-Oriented* causal models introduce into structural causal relationship learning (the primary impact of $\boxed{\text{L2}}$).

2 Relation-Oriented Dimensionality Framework

In the fervent debates surrounding AGI, a pivotal question persists: Can conceptualized symbols and AI systems grounded in symbolism truly represent a human-like comprehension for empirical inquiries Newell (2007); Pavlick (2023). We suggest that the essence lies in enabling informative symbolization. Within our cognitive framework, elements that are abstractly meaningful can be represented as unobservable variables. By directing the learning objective towards extracting these elements from observable data, the symbolized computation within the model remains informative. Such captured relational information has the potential to mirror our cognitive processes of logical deduction.

By Definitions 1 and 2, symbolizing a directional relationship in modeling necessitates two types of variables: the observables $\{\mathcal{X}, \mathcal{Y}\}$, and the unobservables θ and ω. As specified, $\mathcal{X}$ and $\mathcal{Y}$ include both *observational* and *temporal* features. In response, we adopt the concept of a *hyper-dimension* to integrate these unobservable features. Consequently, we establish a framework, as illustrated in Figure 3, to represent relationships as joint distributions across three distinct types of dimensions. For clarity, "feature" refers to the potential variable fully representing a certain distribution of interest.

Figure 3 aims to decompose our cognitive space where relational knowledge is stored. The hyper-dimensional space $\mathbb{R}^H$ is constructed by aggregating all ***unobservable*** relations in our knowledge, such as $(\theta, \omega) \in \mathbb{R}^H$. Conversely, the observational-temporal joint space, $\mathbb{R}^O \cup \mathbb{R}^T$, is considered as the ***observable*** space. In both $\mathbb{R}^O$ and $\mathbb{R}^T$, a temporal dimension consistently signifies the evolution of timing but represents distinct concepts, as outlined in section 2.1. Within such a dimension, linear and nonlinear distributions correspond to *static* and *dynamical* features, respectively, a distinction further explained in section 2.2.

Definition 3. The *Relationship Symbolization* within the proposed Dimensionality Framework.
For the relationship $\mathcal{X} \xrightarrow{\vartheta} \mathcal{Y}$, where $\{\mathcal{X}, \mathcal{Y}\} \in \mathbb{R}^O$ and $\vartheta \in \mathbb{R}^T \cup \mathbb{R}^H$, the ***structuralized*** *relation* ϑ can be decomposed as:

$$\vartheta = \overrightarrow{\theta^1 \dots \theta^T}, \text{ where } (\theta^i, \theta^j) \in \mathbb{R}^H \text{ for any } i \neq j \in \{1, \dots, T\}. \text{ Accordingly,}$$

$$(\vartheta, \omega) = \begin{pmatrix} \vartheta_o \\ \vartheta_\omega \end{pmatrix} = \begin{pmatrix} \theta_o^1 & \dots & \theta_o^T \\ \theta_\omega^1 & \dots & \theta_\omega^T \end{pmatrix} \text{ with any } (\theta_o^i, \theta_o^j) \in \mathbb{R}^H \text{ and } (\theta_\omega^i, \theta_\omega^j) \in \mathbb{R}^H.$$

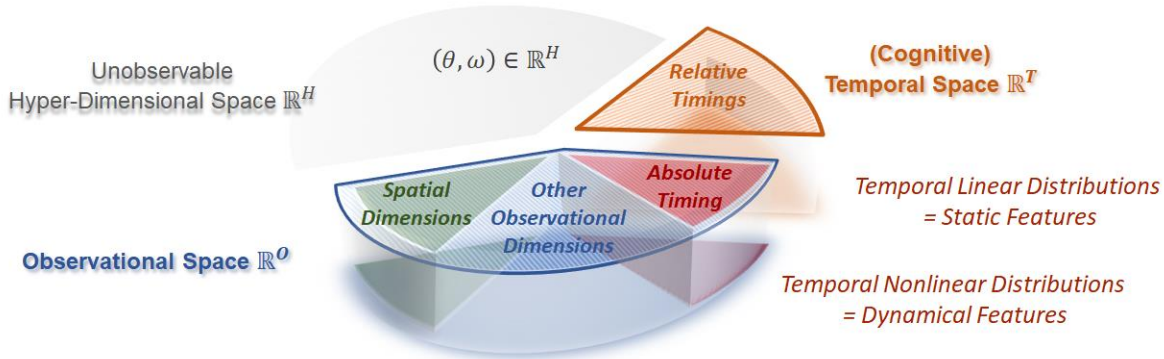


Figure 3: *Relation-Oriented* Dimensionality Framework: splitting the knowledge-storing cognitive space by their accommodated features, where $\{X^t, Y^t\} \in \mathbb{R}^{O-1}$, $\{\mathcal{X}, \mathcal{Y}\} \in \mathbb{R}^O$, and $(\mathcal{X}, \mathcal{Y} \mid \vartheta) \in \mathbb{R}^O \cup \mathbb{R}^T$.

2.1 Absolute Timing vs. Relative Timings

In time series data, the attribute recording observed timestamps, denoted by t, typically reflect the *absolute* timing of reality. However, from a modeling standpoint, the t values are indistinguishable from values of other attributes. Therefore, in Figure 3, the absolute timing $\mathbf{t}$ serves as a standard dimension within the observational space $\mathbb{R}^O$, along which, both $\mathcal{X}$ and $\mathcal{Y}$ are invariably observed as data sequences X^t and Y^t.

In our cognitive space, ***relative*** timings inherently exist Wulf (1994), supporting structuralized relationships beyond the mere ***absolute*** timing $\mathbf{t}$. We thereby designate a distinct "temporal space" $\mathbb{R}^T$, composed of T relative timings as axes, forming T cognitive *timelines*, to integrate temporal features aligned with our relational knowledge Shea (2001). Instead of treating $\{\mathcal{X}, \mathcal{Y}\} \in \mathbb{R}^O$ as individual variables, under ϑ (as per Definition 3), they are jointly distributed across $\mathbb{R}^O$ and $\mathbb{R}^T$, represented as $(\mathcal{X}, \mathcal{Y} \mid \vartheta) \in \mathbb{R}^O \cup \mathbb{R}^T$.

ϑ can span up to T timing dimensions in $\mathbb{R}^T$, with the effect $\mathcal{Y} = \sum_{i=1}^{T} \hat{\mathcal{Y}}^i$ decomposed into T components, each residing in a distinct timing. Crucially, defining ϑ as a "structuralized" relation not only recognizes its multi-dimensionality but also highlights the potential distributional dependence (typically nonlinear) among these dimensions, represented by $(\theta^i, \theta^j) \in \mathbb{R}^H$ in Definition 3, equating to $(\hat{\mathcal{Y}}^i, \hat{\mathcal{Y}}^j) \in \mathbb{R}^O \cup \mathbb{R}^T$. We term these nonlinear temporal dependences as ***dynamical interactions*** for clarity, which necessitate the establishment of $\mathbb{R}^T$ as a distinct "space", rather than mere temporal dimensions within $\mathbb{R}^O$.

For instance, patients' vital signs are recorded daily in a hospital with *absolute* chronological timestamps. However, to assess a medical intervention $\mathcal{Y}$, a uniform series of post-medication events must be selected, for example, spanning from the day after medication to the 30th day. This creates a timeline represented by the axis ticked as $[1, 30]$ to denote the *relative* timing, regardless of *absolute* timestamps of the selected records. Yet, if the intervention involves two distinct aspects, such as the primary effect $\hat{\mathcal{Y}}^1$ and the side effect $\hat{\mathcal{Y}}^2$, and their mutual influences are of interest, then two separate relative timings, $\mathbf{t_1}$ and $\mathbf{t_2}$, must be considered for their individual evolutions, even though both may be labeled as $[1, 30]$.

Remark 2. Although $\mathcal{Y} \in \mathbb{R}^O$ is invariably observed as a sequence along the *absolute* timing $\mathbf{t}$, it may represent an *underlying structure* determined by $\mathcal{X} \xrightarrow{\vartheta} \mathcal{Y}$, spinning the multi-dimensional $\mathbb{R}^T$.

Conventionally, the concept of "temporal dimension" is often simplified as the single absolute timing $\mathbf{t}$, evident from the traditional "spatial-temporal" analysis Alkon (1988); Turner (1990); Andrienko (2003), to recent advancements in language models Wes (2023). However, as emphasized in Remark 2, our cognitive perception of "time" is more complex, which forms the foundation of our causal knowledge Coulson (2009).

To provide an intuitive insight into the implications of neglecting $\mathbb{R}^T$, let's consider an analogy: Imagine ants dwelling on a floor's two-dimensional plane. To predict risks, the scientists among them create two-dimensional models and instinctively adopt the nearest tree as a height reference. They noticed increased disruptions at the tree's first branch, which indeed correlates to the children's heights, given their curiosity. However, without understanding humans as three-dimensional beings, they can only interpret it by adhering to the first branch. One day, after relocating to another tree with a lower height, the ants found the risk presenting at the second branch instead, making their model ineffective. They may conclude that human behaviors are too complex, highlighting the model generalizability issue.

As three-dimensional beings, we inherently lack the capacity to fully integrate the fourth dimension - time - into visual perception. Instead, we conceptualize "space" in three dimensions to incorporate features of the temporal dimension along a timeline within the space, analogous to our "tree". Yet, ants do not need to fully comprehend the three-dimensional world to build a generalizable model; instead, they need only recognize the existence of the "forest", which consists of all "possible trees" with relatively different branch locations. Similarly, in our modeling, we must include the $\mathbb{R}^T$ space, composed of all potential relative timings within our knowledge, although they cannot be directly observed.

Remark 3. *Counterfactuals* can be viewed as posterior distributions in $\mathbb{R}^O \cup \mathbb{R}^T$ given priors in $\mathbb{R}^O$.

The ability to address counterfactual queries, such as "what effect would be if the cause were changed", is a crucial aspect of causal models' significance Scholkopf (2021). A separate cognitive $\mathbb{R}^T$ space allows a more intuitive interpretation of counterfactuals as conditional distributions, which might offer valuable insights in fields like quantum computing. In particular, the observed prior conditions can be viewed as features within $\mathbb{R}^O$; then, all subsequent possibilities can be collectively considered as a distribution across $\mathbb{R}^T$.

2.2 Dynamical vs. Sequential Static

The distributions along a dimension can be broadly classified into *linear* and *nonlinear* categories. Within the temporal dimension, these correspond to ***static*** and ***dynamical*** temporal features, respectively, and can be represented by corresponding variables. Static features are typically linked to specific timestamps. For instance, consider the statement "rain leads to wet floors"; here "wet floors" represents a state that can be identified at a particular point in time. Therefore, it can be denoted as a static variable X_t with a specified timestamp t. In contrast, the expression "floors becoming progressively wetter" necessitates a representation that captures the temporal distribution, to account for changes over time, like $X^t = X_1, \ldots, X_t$. However, this raises the question: Is X^t a dynamical variable or a sequence of static variables?

Within the current machine learning paradigm, the distinction between "static" and "dynamical" is typically made between "models" instead of "variables" PGMadhavan (2016), which refers to whether time is a factor in the model's equations. However, this essentially requires the function $f(X^t; \theta)$ to represent the *dynamics of effect*, inherently encompassed by $\mathcal{Y}$. As a result, the model selection for $f(;\theta)$, as well as the identification of a *static* outcome Y_{t+1}, become crucial in determining how much effect dynamics can be captured Weinberger & Allen (2022), or potentially neglected, which will be discussed in detail in Section 3.

Definition 4. A *Dynamical Outcome* $\mathcal{Y}$ compared to a sequential static Y^τ.
As a *dynamical* variable, $\mathcal{Y} = \langle Y, \tau \rangle \in \mathbb{R}^O$ permits ***nonlinear*** *computational freedom* over τ, whereas a *sequential static* variable $Y^\tau \in \mathbb{R}^{O-1}$ assumes i.i.d or ***linear*** changes along τ. They samely appear to be sequential instances $y^\tau = y_1, \ldots, y_\tau$, while the dynamical significance of $\mathcal{Y}$ is model-dependent.

Definition 4 is based on the proposed *Relation-Oriented* paradigm, wherein the relation $\theta \in \mathbb{R}^H$ and the outcome $\mathcal{Y} \in \mathbb{R}^O \cup \mathbb{R}^T$ are considered individually. Here, θ represents certain unobservable information

within $\mathbb{R}^H$, lacking an explicit distributional representation. This allows $\mathcal{Y}$ to be an individual variable that encompasses the dynamical effects caused by $\mathcal{X}$. Similarly, the cause $\mathcal{X} = \langle X, t \rangle \in \mathbb{R}^O$ can also be a dynamical variable, depending on specific models. For example, RNN models typically formulate $Y_{t+1} = f(\mathcal{X}; \theta)$ with a dynamical cause represented by latent space features, but remaining the outcome static.

Accordingly, the statement "floors becoming progressively wetter" can be roughly considered as "linearly increasing from 0% to 100% in 10 minutes" to be a sequential static feature. It can also be depicted as a continuous nonlinear distribution, a dynamical feature for finer granularity. The latter can cover variances in the former, such as varying progression speeds, which the former cannot. In essence, implementing dynamical variables is crucial for achieving model generalizability across temporal dimensions.

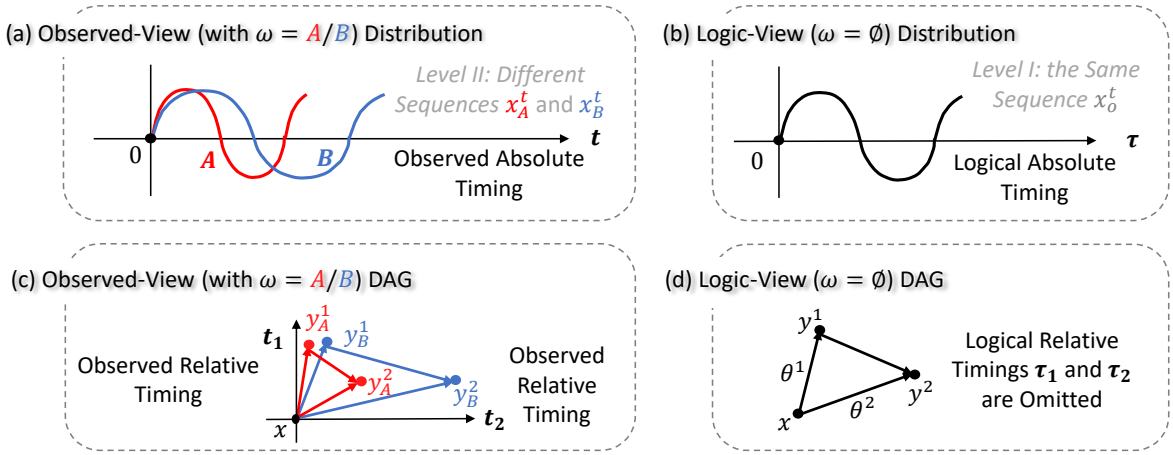


Figure 4: Comparisons of the individualized dynamics from the model's Observed-View, and the generalized dynamics from humans' Logic-View. In (c) and (d), the structuralized relation $\vartheta = \overrightarrow{\theta^1\theta^2}$ with $T = 2$.

When considering a structuralized relation ϑ, a valid generalization process requires the model to remain effective over all temporal dimensions at any level, no matter for the absolute timing within $\mathbb{R}^O$, or the relative timings in cognitive $\mathbb{R}^T$. In our cognition, the generalized causal knowledge ($\omega = \varnothing$) can be instinctively extracted from individualized varied scenarios (with varying ω values). However, the undetectability of (ϑ, ω) implies our models cannot autonomously fulfill this process, irrespective of whether they are AI-based.

Figure 4 showcases models' and humans' perspectives, distinguished as the "Observed-View" and "Logic-View". (a) and (b) compare a simple dynamical distribution within $\mathbb{R}^O$, while (c) and (d) display a DAG structure across two relative timings in $\mathbb{R}^T$, which exhibits a typical *dynamical confounding* scenario. In (c), the static instances y_A^1 and y_B^1 indicate that the two individualized dynamical effects $\mathcal{Y}_A$ and $\mathcal{Y}_B$ reach the same status value y^1 in dimension $\mathbf{t_1}$, signifying that they attain an equivalent magnitude; this is similarly observed in another timing dimension $\mathbf{t_2}$. Notably, the influence from y^1 to y^2 highlights the *dynamical interaction* between the effect components $\hat{\mathcal{Y}}^1$ and $\hat{\mathcal{Y}}^2$, rather than a mere static influence.

Definition 5. The *Dynamical Interaction Confounding* Phenomenon.

For relationship $\mathcal{X} \xrightarrow{\vartheta} \mathcal{Y}$, when effect $\mathcal{Y}$ encompasses multiple dynamics over distinct relative timings, *interactions* among them can lead to *dynamical confounding* within the structuralized $\mathcal{Y}$.

2.3 Informative Hyper-Dimensional Space

In summary, the human-like causal reasoning can be represented as $(\vartheta, \omega) \in \mathbb{R}^T \cup \mathbb{R}^H$. Accordingly, AGI that meets our expectations should adequately encapsulate informative ϑ and ω. Here, $\vartheta \in \mathbb{R}^T \cup \mathbb{R}^H$ denotes the structuralized causality within our knowledge, while $\omega \in \mathbb{R}^H$ indicates the ability to capture nonlinearities in all dimensions (including temporal dynamics), to achieve model generalizability.

Figure 5 provides a fundamental overview of the prevailing relationship modeling methods, highlighting their primary limitations, as briefly summarized in Figure 1. Within this context, ϑ_ω is used to represent generalizable causal structures in AGI, and we identify the two major obstacles in our pursuit of it.

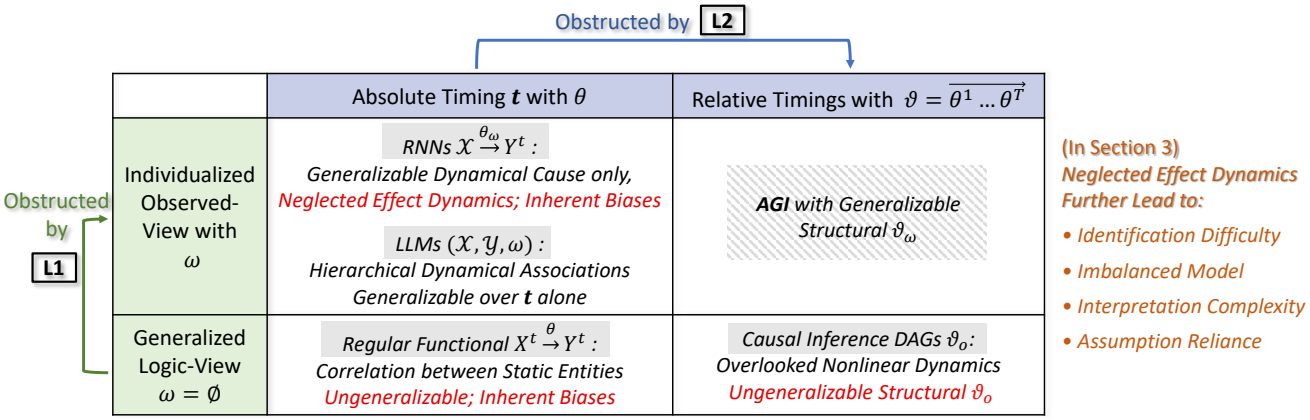


Figure 5: Overview of major obstacles toward AGI (referring to Figure 1). $\boxed{\text{L1}}$ = Undetectable hierarchy by ω that requires dynamical generalizability. $\boxed{\text{L2}}$ = Overlooked multi-dimensional $\mathbb{R}^T$ with relative timings.

Regular causal models derive the functional parameter θ from the correlation between cause and effect events that are priorly identified by absolute timestamps. Notably, Granger causality Granger (1993), a method well-regarded in economics Maziarz (2015), introduces separate temporal sequences for cause (X^t) and effect (Y^τ), suggesting the allowance for multiple timings. However, the significance of recognizing temporal dimensions lies in capturing their featured dynamical evolutions. Without nonlinear computations, distinguishing between $\mathbf{t}$ and τ for static sequential timestamps offers limited meaning.

Similarly, traditional causal inference without accounting for dynamics often neglects explicit relative-timing axes in causal DAGs, as depicted in Figure 4 (d). While inherently adopting a *Relation-Oriented* perspective based on the Logic-View knowledge (ϑ_o), it tends to overlook the Observed-View with varied ω, thus failing to visualize the model generalization needs. To address this, we suggest enhancing conventional DAGs to illustrate dynamical variations across relative timings, as further detailed in Section 4.

Unsurprisingly, AI-based RNNs are increasingly favored in modern relationship learning Xu et al. (2020), considering their proficiency in handling nonlinear causes. RNNs transform the observational sequence X^t into a feature representation in latent space, enabling nonlinear computation over $\mathbf{t}$ to effectively implement dynamical $\mathcal{X}$. However, potential dynamics of the effect $\mathcal{Y}$ are often overlooked, resulting in *imbalanced* causal function $Y_{t+1} = f(\mathcal{X}; \theta)$ with a static outcome Y_{t+1}. This accordingly motivates the emerging trend in *inverse learning* methods Arora (2021). Further details will be discussed in Section 3.

On the other hand, large language models (LLMs) have facilitated the autonomous identification of dynamics under various conditions (ω values) along a timeline that indicates the order of phrases in the semantic space Wes (2023). However, "multiple levels of dynamics" do not necessarily equate to "multiple relative timings". Mutually independent dynamics may represent different temporal dimensions, yet they can still be identified simultaneously from the same absolute timing. In contrast, the $\mathbb{R}^T$ space, comprising relative timings, is capable of accommodating dynamical interactions among these dimensions, as represented by the unobservable association $(\theta^i, \theta^j) \in \mathbb{R}^H$ for any distinct $i, j \in \{1, \ldots, T\}$.

Remark 4. Mere *Temporal Dimensions* could suggest nonlinear independence, while distinct *Relative Timings* imply potential nonlinear dependence, i.e., *dynamical interactions*. The former exists solely along the absolute timing within $\mathbb{R}^O$, whereas the latter together form the multi-dimensional $\mathbb{R}^T$ space.

In essence, current LLMs mainly focus on semantic correlations θ within the absolute timing $\mathbf{t}$ of $\mathbb{R}^O$. Given the consistent sequential semantics of words, omitting relative timings seems justifiable for primarily context-associative learning tasks today. Instead of explicitly extracting θ, most language models implicitly reflect it

by capturing the association $(\mathcal{X}, \mathcal{Y})$. This might contribute to AI's ability to generate intelligent responses without truly "understanding" in a human sense, due to the absence of an informatively extracted θ.

Integrating meta-learning with LLMs could potentially enhance the associative model generalizability Lake (2023), solely on the $\mathbf{t}$ timing. Particularly, given meta-learning's adaptability to diverse conditions in solely observational tasks Hospedales et al. (2021), its application could lead to improved hierarchical $(\mathcal{X}, \mathcal{Y}, \omega)$ to better reflect hierarchical (θ, ω). Consequently, at this stage, given our goal of achieving informative structural knowledge as represented by ϑ_ω, discussing AGI within the current LLM framework might still be premature. However, enabling *Relation-Oriented* meta-learning could potentially bring us closer.

3 Neglected Effect Dynamics in Causality

Traditional causal inference often highlights the interpretability of causal models, notably to be distinguished from mere correlations. In essence, these distinctions are not inherently embedded in the modeling context but are mainly evident in model interpretations, which can potentially guide further causally meaningful improvements for the model. Given the statistical basis of causal inference, the significance of nonlinear temporal dynamics has not been fully embraced yet. This section concentrates on these often overlooked dynamics, aiming to provide a more intuitive understanding of causal learning.

Definition 6. Causality vs. Correlation in the modeling context.

- Causality $\mathcal{X} \xrightarrow{\vartheta} \mathcal{Y}$ is the relationship neccessitating dynamical $\mathcal{X}$ and $\mathcal{Y}$, especially the effect $\mathcal{Y}$.
- Correlation $X^t \xrightarrow{\theta} Y^t$ only requires static cause and effect, possibly sequential static X^t and Y^t.

Timestamp t was first introduced by the Picard-Lindelof theorem in the 1890s, initiating the functional form $Y_{t+1} = f(X_t)$ to represent time evolution. Subsequently, the time series learning methods, like autoregressive models Hyvärinen (2010), facilitate the form of $Y_{t+1} = f(X^t)$ with a sequential causal variable X^t, where the time progress from t to $t{+}1$ is predetermined. For RNNs, the latent space optimization over the representation of X^t is driven by predicting the observed Y_{t+1} through the parameterized relation θ. Consequently, the significant temporal nonlinearity within X^t over $\mathbf{t}$ can be captured, enabling the form of $Y_{t+1} = f(\mathcal{X}; \theta)$ with a dynamical cause $\mathcal{X}$. However, the effect Y_{t+1} remains static, leaving its potentially significant dynamics completely managed by the function f. While f can be selected as nonlinear to enable $\mathcal{X}$, the time evolution from X^t to Y_{t+1} is always left as ***linear***, resulting in *static outcome* sequence $Y^t = Y_1, \ldots, Y_t$.

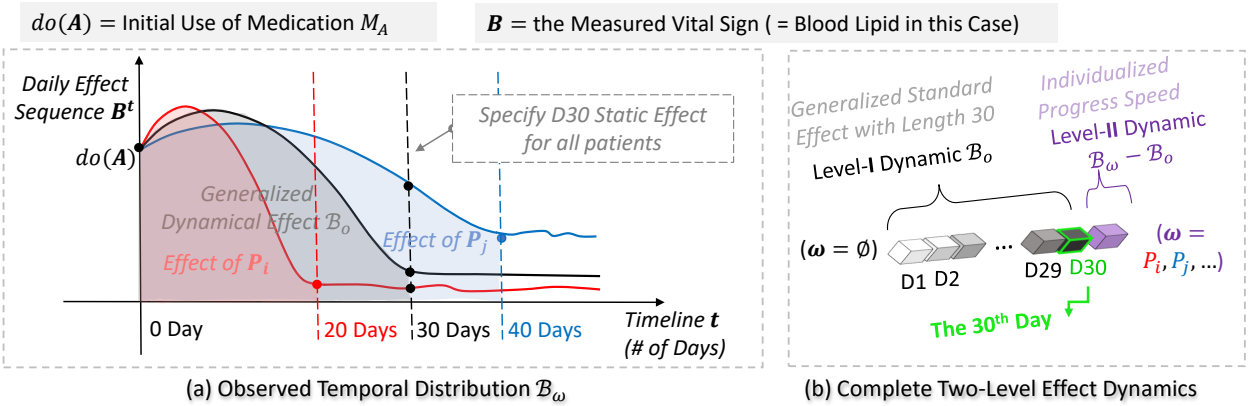


Figure 6: $do(A)$ denotes the initial use of medication M_A for reducing blood lipid B. The goal is to estimate the generalized effect of M_A, i.e., $\mathcal{B}_o$. By the rule of thumb, $\mathcal{B}_o$ needs around 30 days to fully release ($t = 30$ at the black curve elbow). Patient P_i and P_j achieve the same static effect by 20 and 40 days instead.

Figure 6(a) illustrates the often overlooked effect dynamics in traditional causal models. The action $do(A)$ causes dynamical $\mathcal{B}_\omega$ (observed as sequence B^t), disentangled by two levels in (b): Level **I**, the generalized standard sequence $\mathcal{B}_o$ of length 30; Level **II**, the individualized variations $\mathcal{B}_\omega - \mathcal{B}_o$. Assume the unobserved individualized characteristics linearly impact $\mathcal{B}_o$, making $\omega = P_i, P_j, \ldots$ simply represent speeds.

A typical clinical model, like $B_{t+30} = f(do(A_t))$ that averages all patients' D30 static effects as the outcome, turns to neglect D1-D29 within $\mathcal{B}_o$. However, even adopting a sequential outcome B^t (e.g., Granger causality), it remains challenging to accurately estimate $\mathcal{B}_o$ by linear averaging, not to mention further reaching $\mathcal{B}_\omega$. Particularly, it requires the selected records to meet certain criteria, essentially equal to manually defining the boundary of $\mathcal{B}_o$ by exploring all possible ω values.

Such hierarchical dynamical effects are prevalent in fields like epidemic progression, economic fluctuations, and strategic decision-making. They often rely on similar preprocessing to identify specific levels, such as the group-specific learning methodology Fuller et al. (2007). These approaches have become impractical in AI-based applications and may lead to notable information loss in large-scale structural models.

3.1 Identification Difficulty of Dynamical Effect

In a relationship $\mathcal{X} \xrightarrow{\theta} \mathcal{Y}$, an *Observation-Oriented* model can be ideally formulated as $Y^t = f(\mathcal{X};\theta)$ based on existing knowledge, to derive θ and generate the static sequential estimations $\hat{Y_\theta^t} = \hat{Y}_1, \ldots, \hat{Y}_t$ with high accuracy. Yet, two types of errors may present challenges: the discrepancy between the specified outcome sequence and the targeted dynamical effect $\mid \mathcal{Y} - Y^t \mid$; and the modeling error from predetermined function $f(;\theta)$. They contribute to the difficulty of identifying nonlinear effects $\mathcal{Y}$ Zhang (2012).

Specifically, due to the static sequence Y^t, the task of representing neglected dynamics of $\mathcal{Y}$ shifts either to $f(;\theta)$ or to $\mathcal{X}$. In the former scenario, a factor σ representing "disturbance" is integrated into the function, resulting in $f(;\theta+\sigma)$ Zhang (2012). In the latter case, as illustrated in do-calculus Pearl (2012); Huang (2012), the dynamics of $\mathcal{X}$ need to be manually discretized as identifiable temporal events to ensure their observational effects. This enables a fluid transformation from dynamical cause to observational effect, but the identifiability relies on non-experimental data (controllable θ) and can introduce additional complexities.

Considering the *differential* essence of do-calculus, we provide a streamlined reinterpretation of its three core rules from an *integral* viewpoint. Let $do(x_t) = (x_t, x_{t+1})$ indicate the occurrence of an instantaneous event $do(x)$ at time t, with the time step Δt appropriate to ensure the *interventional* effect of $do(x_t)$ identifiable as a function of the resultant distribution at $t+1$. Meanwhile, a separate *observational* effect is provoked by the static x_t. Then, the dynamical cause $\mathcal{X}$ can be discretized as below:

Given $\mathcal{X} \xrightarrow{\theta} Y$, where $\mathcal{X} = \langle X, t\rangle \in \mathbb{R}^{d+1}$ with the augmented $\mathbf{t}$ dimension residing a l-length sequence,

$$\mathcal{X} = \int_0^l do(x_t)\cdot x_t\ dt \quad \text{with} \quad \begin{cases} (do(x_t)=1) \mid \theta, & \textit{Observational} \text{ only (Rule 1)} \\ (x_t=1) \mid \theta, & \textit{Interventional} \text{ only (Rule 2)} \\ (do(x_t)=0) \mid \theta, & \text{No } \textit{interventional} \text{ (Rule 3)} \\ \text{otherwise} & \text{Associated } \textit{observational} \text{ and } \textit{interventional} \end{cases}$$

The effect of $\mathcal{X}$ can be derived as $f(\mathcal{X}) = \int_0^l f_t\big(do(x_t)\cdot x_t\big)\ dt = \sum_{t=0}^{l-1}(y_{t+1} - y_t) = y_l - y_0$

Based on a controllable θ, it addresses three criteria that can preserve conditional independence between *observational* and *interventional* effects, completing the chain rule, but sidesteps more generalized cases. If oppositely defining $\mathcal{Y} = \langle Y, \tau\rangle$ as a dynamical effect, discretizing the dynamics in $do(y)$ remains necessary.

3.2 Imbalance between Cause and Effect

For the model itself, causal directionality (i.e., the roles of cause and effect) may not impose restrictions, although it is often emphasized in model interpretations. Specifically, when selecting a model for a directional relationship $X \to Y$, one could use $Y = f(X;\theta)$ to predict the effect Y, or $X = g(Y;\phi)$ to inversely infer the cause X. Both parameters, θ and ϕ, are obtained from the joint probability $\mathbf{P}(X,Y)$ without imposing modeling constraints. We refer to this as *symmetric directionality* for clarity.

The empirical concerns for modeling directions mainly arise for two reasons: 1) to comply with our intuitive understanding of temporal progression; 2) the current causal modeling exhibits an ***imbalance*** in capturing dynamics between the cause and the effect, with a typical example as RNNs, as represented by $Y = f(\mathcal{X};\theta)$.

Given the symmetric directionality, and to capitalize on the imbalance, inverse learning methodology Arora (2021) has recently garnered increasing attention, to achieve autonomous dynamical effect identification by inversely assigning the effect as the cause within RNNs. However, this approach is unsuitable to address the structuralized relation ϑ. Specifically, the overlooked relative timings within ϑ and the neglected dynamical interactions among them can introduce inherent bias into the effect, which the inverse of the model cannot eliminate. This issue is further detailed in Section 4.

Another factor contributing to the imbalance is the increased difficulty when specifying effect sequence Y^t compared to cause sequence X^t. While organizing sequential data around a major causal event (e.g., days of heavy rain) is feasible, pinpointing the precise onset of subsequent effects (e.g., the exact day a flood began due to the rain) remains a more complex task.

Remark 5. By indexing through θ, simultaneous optimization of $\mathcal{X}$ and $\mathcal{Y}$ can be achieved, mitigating their imbalance and enabling autonomous identification of both dynamical variables.

The *Relation-Oriented* modeling approach seeks to autonomously derive θ from the feature representations of $\mathcal{X}$ and $\mathcal{Y}$ within the latent space. Specifically, the initial sequences X^t and Y^t are transformed into a latent space, $\mathbb{R}^L$, which allows nonlinear computational freedom in their temporal dimensions. Then, a neural network representing θ can be trained between them without relying on prior assumptions.

The training process uses $\mathcal{X}$ as the input and $\mathcal{Y}$ as the output, indexed through θ, facilitating the concurrent optimization of both dynamical representations. Consequently, this yields sequentially associated $(\mathcal{X}, \theta, \hat{\mathcal{Y}}_\theta)$, with each individual representation maintaining. The implementation will be outlined in Chapter II.

3.3 Interpretation Complexity

Since effect dynamics are often partially overlooked, traditional causal inference introduces the concept of "hidden confounder" to enhance model interpretability. For example, the node E in Figure 7 (a) symbolizes the unobserved individualized characteristics in the scenario depicted in Figure 6.

However, this approach does not necessarily require collecting additional data to identify E. This might lead to an illogical implication: "Our model is biased due to some unknown factors we don't intend to explore." Indeed, this strategy employs a solely observational causal variable, E, to account for the overlooked dynamical effect features. While E remains unknown, its inclusion can complete the model interpretation. Yet, from the modeling perspective, as illustrated in Figure 7(b), the associative cause $do(A) * E$ remains unknown, failing to provide a modelable relationship for addressing $(\theta, \omega) = \begin{pmatrix} \theta_o \\ \theta_\omega \end{pmatrix}$.

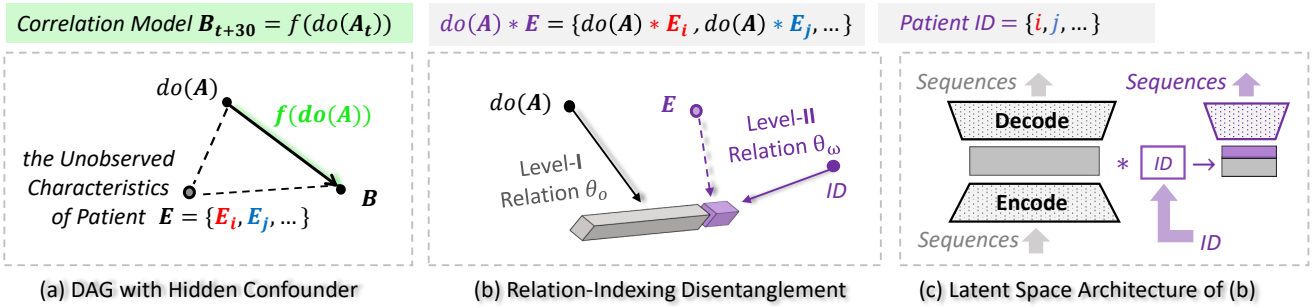
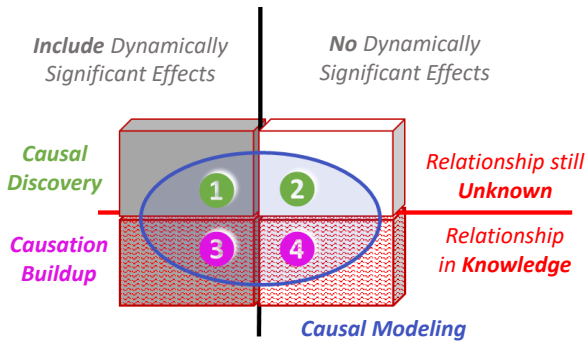


Figure 7: (a) Traditional causal inference DAG. (b) Hierarchical disentanglement of effect dynamics through relation-indexing. (c) Autoencoder-based generalized and individualized reconstruction processes.

Fundamentally, incorporating a hidden confounder can improve the model's interpretability but not its generalizability. In contrast, the *Relation-Oriented* approach does not require extra modeling; it leverages θ as indices to extract $\hat{\mathcal{Y}}_\theta$, enabling the use of any observed identifier associated with ω, such as patient IDs. As illustrated in (c), this hierarchical disentanglement of representations in the latent space can effectively encapsulate effect dynamics.

3.4 Causal Assumptions Reliance

Due to the frequently overlooked effect dynamics, traditional causal learning typically relies on foundational causal assumptions to validate practical applications. In Figure 8, we categorize causal model applications into four distinct scenarios based on two aspects: Firstly, depending on whether the predetermination for θ is based on knowledge, they are divided into Causal Discovery and Causation Buildup. Secondly, they are further differentiated by the dynamical significance of their effects.



	Modeled Relation	Modeled Directionality
1	Observational Only. Undiscovered Dynamics covered by *Faithfulness Assumption.*	Observational Data Determined. Not Logically Meaningful.
2	Observational Only. Aligned with Knowledge.	Observational Data Determined. Maybe Logically Suggestive.
3	Knowledge Determined. Unmodeled Dynamics covered by *Hidden Confounders* or *Sufficiency Assumption.*	Knowledge Determined.
4	Knowledge Determined.	Knowledge Determined.

Figure 8: Categories of currently prevalent causal learning applications. The left rectangular cube illustrates causally meaningful relationships in logic, with the potentially modelable scope highlighted in blue.

As depicted in Figures 6 and 7, the individualized dynamical features are easily overlooked in a generalized causation buildup process. Based on existing knowledge, some unobserved entities may be identified as hidden confounders, thereby enriching model interpretations. Nonetheless, if such identification is not easy, the foundational *Causal Sufficiency* assumption may lead to the complete neglect of these dynamics, presuming that all potential "hidden confounders" have been observed in the system.

On the other hand, causal discovery typically unearths structural relationships by detecting dependence among observables, but is usually confined to their observational attributes excluding temporal features. If their dynamical features are not crucial, discovered associations can provide valuable insights into the underlying correlations; if they are essential, significant dynamics might be overlooked due to the *Causal Faithfulness* assumption, which suggests that captured observables can fully represent the causal reality.

Furthermore, although the discovered relationships are directional, these directions frequently lack a logical causal implication. Consider X and Y with predetermined directional models $Y = f(X;\theta)$ and $X = g(Y;\phi)$. The direction $X \to Y$ would be favored if $\mathcal{L}(\hat{\theta}) > \mathcal{L}(\hat{\phi})$. Let $\mathcal{I}_{X,Y}(\theta)$ denote the Fisher information about θ given $\mathbf{P}(X,Y)$. Use $p(\cdot)$ as the density function, and $\int_X p(x;\theta)dx$ remains constant in this context. Then:

$$\begin{aligned}\mathcal{I}_{X,Y}(\theta) &= \mathbb{E}[(\frac{\partial}{\partial\theta}\log p(X,Y;\theta))^2 \mid \theta] = \int_Y\int_X (\frac{\partial}{\partial\theta}\log p(x,y;\theta))^2 p(x,y;\theta)dxdy \\ &= \alpha\int_Y (\frac{\partial}{\partial\theta}\log p(y;x,\theta))^2 p(y;x,\theta)dy + \beta = \alpha\mathcal{I}_{Y|X}(\theta) + \beta, \text{with } \alpha,\beta \text{ being constants.}\end{aligned}$$

$$\text{Then, } \hat{\theta} = \arg\max_\theta \mathbf{P}(Y \mid X,\theta) = \arg\min_\theta \mathcal{I}_{Y|X}(\theta) = \arg\min_\theta \mathcal{I}_{X,Y}(\theta), \text{ and } \mathcal{L}(\hat{\theta}) \propto 1/\mathcal{I}_{X,Y}(\hat{\theta}).$$

The inferred directionality indicates how informatively the data reflects the two predetermined parameters, often restrictively based on observational dependency. Consequently, such directionality is not logical but could be dominated by the data collection process, with the predominant entity deemed the "cause", consistent with some existing conclusions Reisach (2021); Kaiser (2021). Even when informative θ and ϕ are incorporated based on knowledge, they might not provide insights for dynamically significant causal relations.

4 Relative Timings in Structural Causality

Consider a structural relationship $\mathcal{Y} \xleftarrow{\theta_1} do(X) \xrightarrow{\theta_2} \mathcal{Z}$, where two dynamical effects of $do(X)$ progress along distinct relative timings $\mathbf{t_1}$ and $\mathbf{t_2}$. Initially, $\mathcal{Y}$ and $\mathcal{Z}$ are identified as sequences Y^t and Z^t according to absolute timing $\mathbf{t}$. If they are nonlinearly independent without interaction, denoted as $\theta_1 \perp \theta_2 \in \mathbb{R}^H$, then it may not be challenging for AI models like inverse RNNs to capture $\vartheta = \overrightarrow{\theta_1\theta_2}$ accurately, by utilizing the functional form $do(X) = f\big((Y,Z)^t; \vartheta\big)$ with associative identification $(Y,Z)^t = \big((Y,Z)_1, \ldots, (Y,Z)_t\big)$.

However, if $\mathcal{Y}$ and $\mathcal{Z}$ interact, like from $\mathcal{Y}$ to $\mathcal{Z}$, illustrated as an edge $Y \to Z$, implying $(\theta_1, \theta_2) \in \mathbb{R}^H$, then the associative identification $(Y,Z)^t$ could introduce ***inherent bias*** due to the *dynamical interaction confounding*, thereby reducing the model robustness and generalizability. Instead, it is necessary to initialize Y^t and Z^t individually, and then engage in a two-step *relation-indexed learning* to sequentially obtain $\mathcal{Y} = f_1(do(X); \theta_1)$, and $\mathcal{Z} = f_2(do(X) \mid \mathcal{Y}; \theta_2)$.

This section will first demonstrate the *inherent bias* through an intuitive example (section 4.1), explore its impact on the generalizability of structural causal models (section 4.2), and finally discuss the advancements and challenges on our path toward incorporating structural causal knowledge within AI (section 4.3).

4.1 Scheme of the Inherent Bias

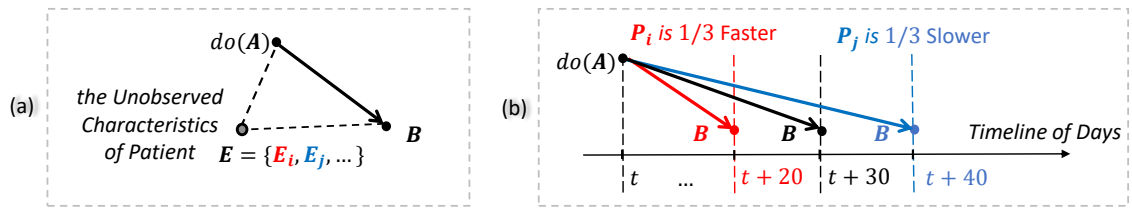


Figure 9: (a) Initial DAG introducing hidden E. (b) Enhanced DAG (Directed Acyclic Graph).

Figure 9(a) revisits the hidden-confounder inclusion depicted in Figure 6. To clearly visualize the dynamical variations across multi-dimensional relative timings, we propose an enhancement to the conventional causal DAGs. This enhancement, as shown in (b), is carried out through two steps:

1. Consider dynamically significant effects and integrate their relative timings as individual axes.
2. Use edge lengths to signify timespans needed for reaching a certain effect magnitude in a static value.

Figure 10(a) depicts a structural relationship $\mathcal{B} \xleftarrow{\theta_1} A \xrightarrow{\theta_2} \mathcal{C}$, extending from the scenario in Figure 9(b), with A succinctly replacing $do(A)$. It features two distinct effects: the primary effect $\mathcal{B}$ via θ_1, represented by the edge $\overrightarrow{AB}$ leading to a static value for vital sign B; and a side effect $\mathcal{C}$ via θ_2 on another vital sign C, indicated by edge $\overrightarrow{AC}$. Notably, $\mathcal{C}$ can influence $\mathcal{B}$, suggesting that A may also indirectly affect B through C, thereby creating *dynamical confounding* among the three entities across two timing axes $\mathbf{t_1}$ and $\mathbf{t_2}$. For simplicity, we fix the timespan of $\overrightarrow{AC}$ at 10 days for all patients, and focus on modeling the static outcome B to predict the average fully-released medical effect in this population.

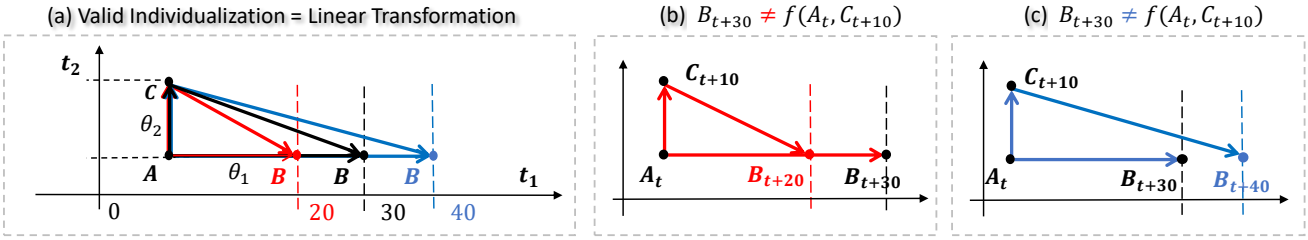


Figure 10: (a) The enhanced DAG with two relative timing axes. (b) (c) Violations of the Markov condition, when timestamps are specified for static effects identification, to construct Structural Causal Models (SCMs).

From a geometrical view, the triangle over nodes $\{A, B, C\}$ should remain closed across all populations and individuals to represent the same relationship, as supported by the *Causal Markov* condition. Accordingly, the generalization (and also individualization) process can be geometrically viewed as a *linear transformation* of the causal DAG, depicted as "stretching" the triangle along $\mathbf{t_1}$ at various ratios, as in Figure 10(a).

In conventional SCMs, the status of B is typically derived by setting an average timespan for the full release of medicine along $\overrightarrow{AB}$, say 30 days in this case. As illustrated in (b) and (c), the SCM function fails to shape a valid DAG for individual patients, represented by P_i in red and P_j in blue. Consequently, sequential biases would be implied when extending to estimate a sequential outcome like $B^t = (B_1, \ldots, B_{30})$.

Definition 7. *Inherent Biases* within Observation-Oriented SCMs.
The *inherent bias* may occur when identifying causal effect if it contains: 1) dynamical interaction confounding across multiple relative timings, and 2) undetectable hierarchy represented by ω.

In this simplified scenario, an inverse RNN model, formulated as $A = f\big((B, C)^t\big)$, could be effective due to the assumed independence. However, it is impractical to assume independence or the absence of confounded dynamics for all effects. This is particularly true in large models dealing with complex causal structures, where inherent biases can accumulate, ultimately jeopardizing the model's robustness.

4.2 Inherently Restricted Generalizability

To address the issues of *confounded dynamics*, traditional causal inference employs various methods to "de-confound"' them, such as cutting off interaction through propensity score matching Benedetto (2018) and backdoor adjustment Pearl (2009). However, these techniques often require specific tailoring for each application, necessitating manual identification of dynamically significant causal effects. Given the black-box nature and large scale of AI models, such manual adjustments have become increasingly impractical.

Moreover, these methods primarily focus on adapting to statistical linear models, which may not effectively contribute to dynamical generalizability. Subsequently, we will use a practical scenario to clearly illustrate how the specification of timestamps for effects inherently hinders the generalizability of the formulated SCMs.

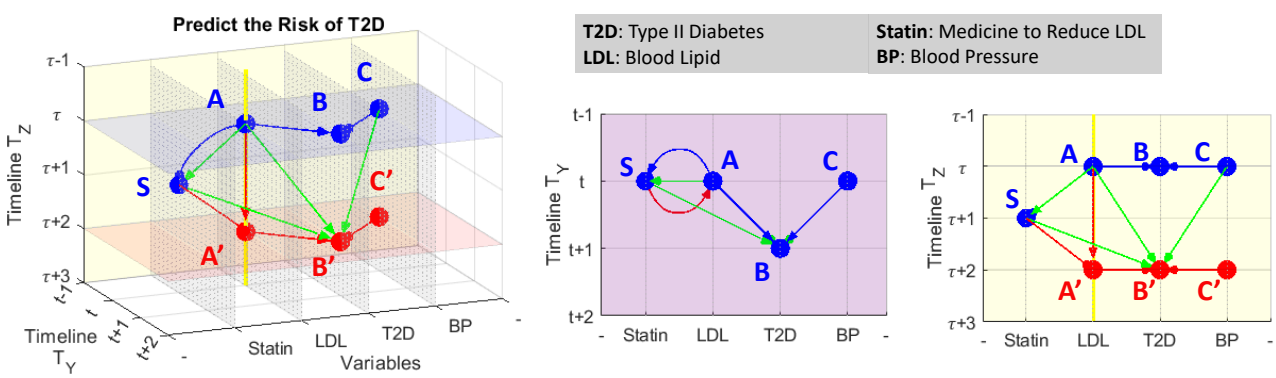


Figure 11: A DAG with two relative timing exes $\mathcal{T}_Y$ and $\mathcal{T}_Z$. The formulated SCM $B' = f(A, C, S)$ evaluates the effect of using S to reduce T2D risks at B'. On $\mathcal{T}_Y$, the step Δt from t to $(t+1)$ allows A and C to fully influence B. The step $\Delta\tau$ on $\mathcal{T}_Z$, from $(\tau+1)$ to $(\tau+2)$, let S fully release to forward status A to A'.

Figure 11 displays an enhanced 3D view DAG, where Δt and $\Delta\tau$ signify actual time spans, particularly within the current population, to support the causal reasoning represented by this structure. Consider the triangle $SA'B'$: As each unit of effect from S delivered to A' (spent $\Delta\tau$), it immediately starts to impact B' through $\overrightarrow{A'B'}$ (Δt needed); meanwhile, the next unit begins generation at S. This dual action runs concurrently until S's effect fully reaches B', representing the single edge $\overrightarrow{SB'}$ within the SCM.

Due to the equation $\overrightarrow{SB'} = \overrightarrow{SA'} + \overrightarrow{A'B'}$, specifying the time span of $\overrightarrow{SB'}$ inherently determines the $\Delta t : \Delta\tau$ ratio based on the current population's performance, thereby fixing the shape of the ASB' triangle in the

DAG space. If we focus solely on the accuracy of the estimated mean effect for this population, the SCM function $B' = f(A, C, S)$ can be effective. However, given that the preset $\Delta t : \Delta\tau$ ratio is not universally applicable, the generalizability of the established SCM to other populations becomes questionable.

4.3 Developments Toward Causal Reasoning AI

In the pursuit of causal reasoning in machine learning, modeling techniques have evolved from capturing mere associations to learning observational correlations, ultimately advancing to structural causality modeling that incorporates the cognitive temporal space $\mathbb{R}^T$. Figure 12 summarizes this evolution in an upward trajectory.

Model	*Principle*	*Cause*	*Relation & Direction*	*Effect*	*Handle Undetectable Hierarchy*	*Capture Dynamics*
Mechanistic or Physical	$\mathcal{Y} = f(\mathcal{X}; \theta)$	Dynamical $\mathcal{X} = \langle X, \boldsymbol{t} \rangle$	by Knowledge	Dynamical $\mathcal{Y} = \langle Y, \boldsymbol{\tau} \rangle$	Yes	Yes
Relation-Indexing Approach	*Given* $\boldsymbol{P}(\mathcal{X}, \mathcal{Y})$ & $\mathcal{X} \xrightarrow{\vartheta} \mathcal{Y}$	Dynamical $\mathcal{X} = \langle X, \boldsymbol{t} \rangle$	by Representation $= f(\mathcal{X}, \vartheta, \hat{\mathcal{Y}}_\vartheta)$	Dynamical $\mathcal{Y} = \langle Y, \boldsymbol{\tau} \rangle$	Yes	Yes
Structural Causal Learning	*Given* $\boldsymbol{P}(X, Y)$ & $X \to Y$ $Y = f(X; \theta)$	Observational Sequence X^t	$X \to Y$ with Predetermined θ	Static Y_τ	?	?
Graphical Causal Discovery	*Given* $\boldsymbol{P}(X, Y)$ *Find* $\mathcal{L}(Y\|X; \theta) > \mathcal{L}(X\|Y; \theta)$	Observational X	Associated (X, Y) with insights into Correlation	Observational Y	?	No
Common Cause Model	*Given* $\boldsymbol{P}(X, Y\|Z)$	Observational X	Conditional Associated $(X, Y\|Z)$	Observational Y	?	No
i.i.d. Associative Model	*Given* $\boldsymbol{P}(X, Y)$	Observational X	None	Observational Y	No	No

Figure 12: Simple taxonomy of models (partially refer to Scholkopf (2021) Table 1), from more data-driven upward to more knowledge-driven. "?" means depending on the practice.

Given AI's ability to learn temporal dynamics, the current challenge is to deal with the dynamical interactions within $\mathbb{R}^O \cup \mathbb{R}^T$. However, as demonstrated in sections 4.1 and 4.2, even when modeling static outcomes, specifying timestamps to accomplish the required temporal events identification can still introduce inherent biases. Consequently, we need a knowledge-aligned manner to capture dynamics, specifically to align with our cognitive understanding, where the abstract relative timings underpin the causal knowledge's structures. This suggests a shift from the current *Observation-Oriented* paradigm towards a new one. Physical models, explicitly incorporated in temporal dimensional computation, may offer valuable insights into this prospect.

Under the observational i.i.d. assumption, initial models only approximate associations, proved unreliable for causal reasoning Pearl et al. (2000); Peters et al. (2017). Subsequently, the common cause principle highlights the significance of the nontrivial condition, to distinguish a relationship from statistical dependencies Dawid (1979); Geiger (1993), providing a basis for constructing graphical models Peters et al. (2014). The initial graphical model relies on conditional dependencies to construct Bayesian networks, with limited causal relevance Scheines (1997). Then, causally significance emphasizes the capability of addressing counterfactual queries Scholkopf (2021), like the structural equation models (SEMs) and functional causal models (FCMs) Glymour et al. (2019); Elwert (2013), which leverage prior knowledge to establish causal structures.

State-of-the-art deep learning on causality encodes the discrete, DAG-structural constraint into continuous optimization functions Zheng et al. (2018; 2020); Lachapelle et al. (2019), enabling advanced efficiency, but without noticeable generalizability, evident from the restricted successes in applications like the neural architecture search (NAS) Luo (2020); Ma (2018). This is reasonable, since the neglected relative timings can lead to inherent biases amplified through complex structures to become significant.

Scholkopf (2021) summarized our confronting key challenges toward generalizable causal-reasoning AI: 1) limited model robustness, 2) insufficient model reusability, and 3) inability to handle data heterogeneity (i.e., undetectable hierarchies). They are intrinsically linked to the demonstrated inherent biases.

Chapter II: Realization of Proposed Relation-Oriented Paradigm

This chapter introduces the proposed *Relation-Indexed Representation Learning* (RIRL) method, a baseline realization of the raised *Relation-Oriented* modeling paradigm. RIRL primarily focuses on autonomously identifying dynamical effects, in the form of relation-indexed representations in the latent space. In the context of structural modeling, RIRL enables hierarchical disentanglement of effects, according to given DAGs, as a manner of realizing dynamical generalizability across undetectable levels within knowledge. As a baseline realization, RIRL is suitable for applications with mature structural causal knowledge, and plenty of data to support neural network training on each known causal relationship.

First, Section 5 details the technique for extracting relation-indexed representations. Then, building on this, Section 6 presents the RIRL method of establishing structural causal models in the latent space. Lastly, Section 7 provides experiments to validate RIRL's efficacy in autonomously identifying effects.

5 Relation-Indexed Representation

In the relationship $\mathcal{X} \to \mathcal{Y}$, we define dynamical $\mathcal{X} = \langle X, t \rangle \in \mathbb{R}^{d+1} \subseteq \mathbb{R}^O$ and $\mathcal{Y} = \langle Y, \tau \rangle \in \mathbb{R}^{b+1} \subseteq \mathbb{R}^O$, given their solely observational variables, $X \in \mathbb{R}^d$ and $Y \in \mathbb{R}^b$. $\mathcal{X}$ is observed as a data sequence, represented by $X^t = X_1, \ldots, X_t$ with a pre-determined length l_x. For clarity, hereafter in this chapter, its instance x^t will be considered as a $(d * l_x)$-dimensional vector, denoted by $\overrightarrow{x}$ (or x for briefty). Similarly, $\mathcal{Y}$ is observed as the data sequence Y^t with a pre-determined length l_y, and its instance is referred to as a $(b * l_y)$-dimensional vector $\overrightarrow{y}$ (or y for briefty).

The relation-indexed representation aims to formulate $(\mathcal{X}, \theta, \hat{\mathcal{Y}}_\theta)$ in the latent space $\mathbb{R}^L$, beginning with an *initialization* to transform X^t and Y^t to be latent space features. For the sake of clarity, we use $\mathcal{H} \in \mathbb{R}^L$ and $\mathcal{V} \in \mathbb{R}^L$ to refer to the latent representations of $\mathcal{X} \in \mathbb{R}^O$ and $\mathcal{Y} \in \mathbb{R}^O$, respectively.

The modeling process is to optimize the neural network function $f(;\theta)$ in $\mathbb{R}^L$, with $\mathcal{H}$ as its input and $\mathcal{V}$ as the output. This process simultaneously refines $\mathcal{H}$, θ, and $\mathcal{V}$, for ultimately achieving $(\mathcal{H}, \theta, \hat{\mathcal{V}}_\theta) = (\mathcal{X}, \theta, \hat{\mathcal{Y}}_\theta)$. The refining will present as the distance minimization between $\mathcal{H}$ and $\mathcal{V}$ within $\mathbb{R}^L$. Consequently, the dimensionality L of the latent feature space must satisfy $L \geq rank(\mathcal{X}, \theta, \mathcal{Y})$, raising a technical challenge that L is larger than the dimensionality of $\overrightarrow{x}$ or $\overrightarrow{y}$.

Remark 6. The variable *initialization* necessitates a *higher-dimensional* representation autoencoder.

5.1 Higher-Dimensional Autoencoder

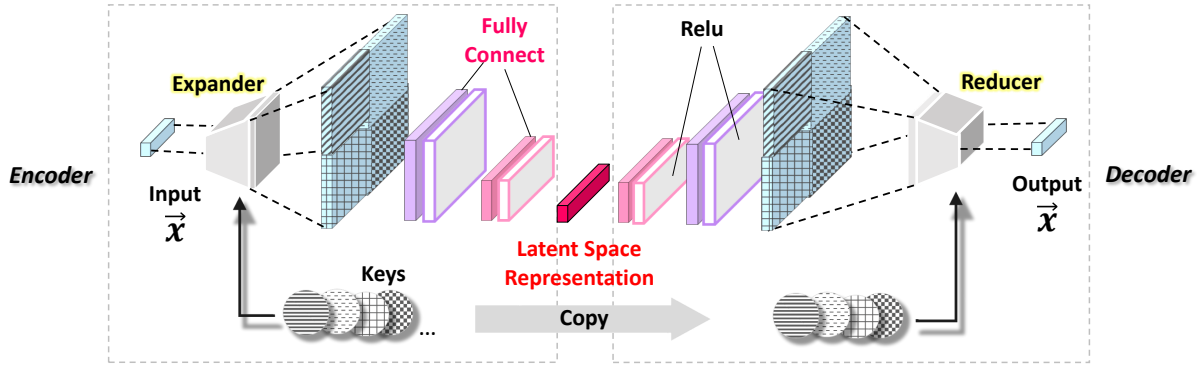


Figure 13: *Invertible* autoencoder architecture for extracting *higher-dimensional* representations.

Autoencoders are commonly used for dimensionality reduction, especially in structural modeling that involves multiple variables Wang (2016). In contrast, RIRL aims to model individual causal relationships sequentially within a higher-dimensional latent space $\mathbb{R}^L$, as to hierarchically construct the entire causal structure. As

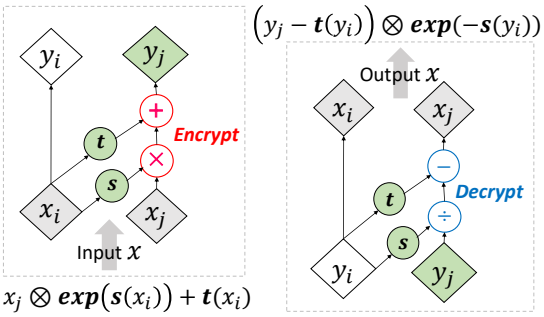


Figure 14: Expander (left) and Reducer (right).

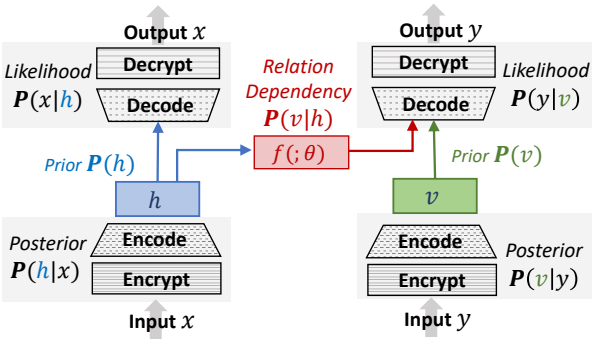


Figure 15: Relationship model architecture.

illustrated in Figure 13, the designed autoencoder architecture is featured by the symmetrical *Expander* and *Reducer* layers (source code is available [1]). The Expander magnifies the input vector $\overrightarrow{x}$ by capturing its higher-order associative features, while the Reducer symmetrically diminishes dimensionality and reverts to its initial state. For precise reconstruction, the *invertibility* of these processes is essential.

The Expander showcased in Figure 13 implements a *double-wise* expansion. Here, every duo of digits from $\overrightarrow{x}$ is encoded into a new digit using an association with a random constant, termed the *Key*. This *Key* is generated by the encoder and replicated by the decoder. Such pairwise processing of $\overrightarrow{x}$ expands its length from $(d*l_x)$ to be $(d*l_x-1)^2$. By leveraging multiple *Keys* and concatenating their resultant vectors, $\overrightarrow{x}$ can be considerably expanded, ready for the subsequent dimensionality-reduced representation extraction. The four blue squares with unique grid patterns represent expansions by four distinct *Keys*, with the grid patterns acting as their "signatures". Each square symbolizes a $(d*l_x-1)^2$ length vector. Similarly, higher-order expansions, like *triple-wise* across three digits, can be achieved with adapted *Keys*.

Figure 14 illustrates the encoding and decoding processes within the Expander and Reducer, targeting the digit pair (x_i, x_j) for $i \neq j \in 1, \ldots, d$. The Expander function is defined as $\eta_\phi(x_i, x_j) = x_j \otimes exp(s(x_i)) + t(x_i)$, which hinges on two elementary functions, $s(\cdot)$ and $t(\cdot)$. The *Key* parameter, ϕ, embodies their weights, $\phi = (w_s, w_t)$. Specifically, the Expander morphs x_j into a new digit y_j utilizing x_i as a chosen attribute. In contrast, the Reducer symmetrically uses the inverse function η_ϕ^{-1}, defined as $(y_j - t(y_i)) \otimes exp(-s(y_i))$.

This approach circumvents the need to compute s^{-1} or t^{-1}, thereby allowing more flexibility for nonlinear transformations through $s(\cdot)$ and $t(\cdot)$. This is inspired by the groundbreaking work in Dinh et al. (2016) on invertible neural network layers employing bijective functions.

5.2 Optimization Steps

Consider instances x and y of $\mathcal{X}$ and $\mathcal{Y}$, with corresponding representations h and v in $\mathbb{R}^L$. The latent dependency $\mathbf{P}(v|h)$ is used to train the relation function $f(;\theta)$, as illustrated in Figure 15. In each iteration, the modeling process undergoes three optimization steps:

1. Optimizing the cause-encoder by $\mathbf{P}(h|x)$, the relation model by $\mathbf{P}(v|h)$, and the effect-decoder by $\mathbf{P}(y|v)$ to reconstruct the relationship $x \to y$, represented as $h \to v$ in $\mathbb{R}^L$.
2. Fine-tuning the effect-encoder $\mathbf{P}(v|y)$ and effect-decoder $\mathbf{P}(y|v)$ to accurately represent y.
3. Fine-tuning the cause-encoder $\mathbf{P}(h|x)$ and cause-decoder $\mathbf{P}(x|h)$ to accurately represent x.

During this process, the values of h and v are iteratively adjusted to reduce their distance in $\mathbb{R}^L$, with $f(;\theta)$ serving as a bridge to span the distance. Here, the hyper-dimensional variable $\theta \in \mathbb{R}^H$ acts as the index, guiding the output of $f(;\theta)$ to encapsulate associated representations $(\mathcal{H}, \theta, \hat{\mathcal{V}}_\theta)$. From $\hat{\mathcal{V}}_\theta$, the effect component $\hat{\mathcal{Y}}_\theta$ can be reconstructed. Within the system, for each effect, a series of such relation functions $\{f(;\theta)\}$ is maintained, indexing diverse levels of causal inputs for sequentially building the structural model.

[1] https://github.com/kflijia/bijective_crossing_functions/blob/main/code_bicross_extracter.py

6 RIRL: Building Structural Models in Latent Space

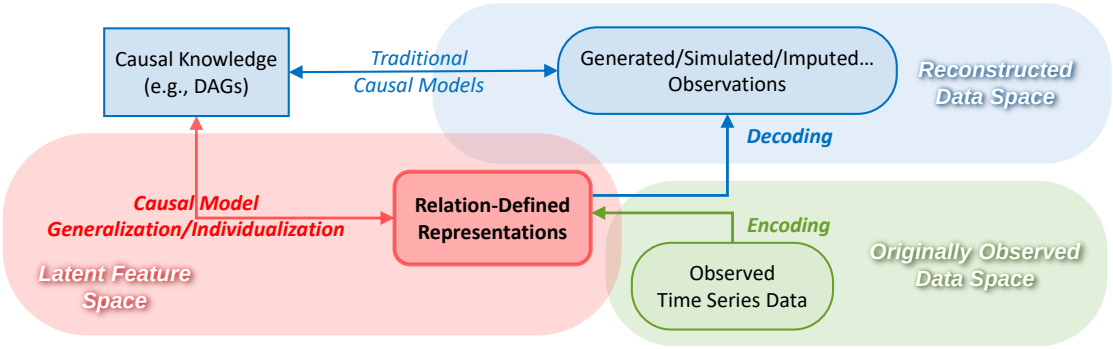


Figure 16: How Relation-Indexed Representation Learning (RIRL) contributes to traditional models.

By sequentially constructing relation-indexed representations for each pairwise relationship within the causal DAG, we can progressively achieve hierarchically disentangled representations for individual nodes. These representations align with the ω levels defined by the global structure, while simultaneously, the entire causal DAG structure will also be established accordingly. Subsequently, section 6.1 introduces the method for stacking higher-level representations upon established lower-level ones for individual effects. Section 6.2 then elaborates on the complete factorization process for hierarchical disentanglement. Finally, section 6.3 discusses a causal discovery algorithm within the latent space for initialized variable representations.

Figure 16 demonstrates how the RIRL method can encapsulate the black-box nature of AI within the latent space while simultaneously generating interpretable observations. This ability can be used to enhance existing *Observation-Oriented* models, for instance, by facilitating on-demand counterfactual simulations. Meanwhile, in the latent space, these cryptic representations, although opaque to human interpretation, play a crucial role in achieving model generalization and individualization. These processes are latently managed by AI and remain exclusive to human comprehension.

6.1 Stacking Hierarchical Representations

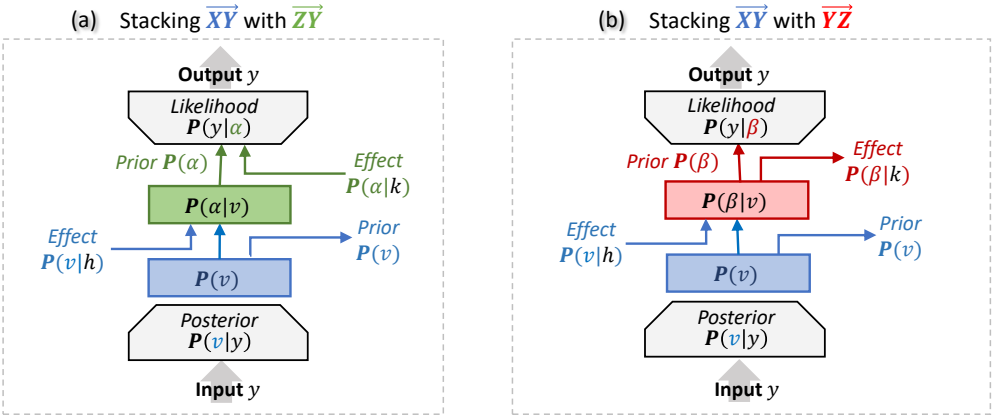


Figure 17: Stacking relation-indexed representations to construct hierarchy.

A structural relationship can be represented by a causal graph, denoted as G. To construct models in the latent space, the latent dimensionality L must be sufficiently large to adequately represent G. Let's denote a data matrix augmented by all observational attributes in G as $\mathbf{X}$. Given the need to include informative relations $\{\theta\}$ for the edges in G, it is essential that $L > rank(\mathbf{X}) + T$, where T indicates the number of dynamically significant variables (i.e., nodes) within G.

The PCA principle posits that the space $\mathbb{R}^L$ learned by the autoencoder is spanned by the top principal components of $\mathbf{X}$ Baldi (1989); Plaut (2018); Wang (2016). Hypothetically, reducing L below $rank(\mathbf{X})$ may yield a less adequate but causally more significant latent space through better alignment of dimensions Jain (2021) (Further exploration in this direction is warranted). Bypassing a deep dive into dimensionality boundaries, we rely on empirical fine-tuning for the experiments in this study (reducing L from 64 to 16).

Consider a causal structural among $\{\mathcal{X}, \mathcal{Y}, \mathcal{Z}\}$, with their corresponding representations $\{\mathcal{H}, \mathcal{V}, \mathcal{K}\} \in \mathbb{R}^L$ initialized by three autoencoders, respectively. Figure 17 illustrates the hierarchical representations buildup. Here, two stacking scenarios are displayed based on varying causal directions. With the established $\mathcal{X} \to \mathcal{Y}$ relationship in $\mathbb{R}^L$, the left-side architecture finalizes the $\mathcal{X} \to \mathcal{Y} \leftarrow \mathcal{Z}$ structure, while the right-side focuses on $\mathcal{X} \to \mathcal{Y} \to \mathcal{Z}$. Through the addition of a representation layer, hierarchical disentanglement is formed, allowing for various input-output combinations (denoted as $\mapsto$) according to specific requirements.

For example, on the left, $\mathbf{P}(v|h) \mapsto \mathbf{P}(\alpha)$ represents the $\mathcal{X} \to \mathcal{Y}$ relationship, whereas $\mathbf{P}(\alpha|k)$ implies $\mathcal{Z} \to \mathcal{Y}$. Conversely, on the right, $\mathbf{P}(v) \mapsto P(\beta|k)$ denotes the $\mathcal{Y} \to \mathcal{Z}$ relationship with $\mathcal{Y}$ as input. Meanwhile, $\mathbf{P}(v|h) \mapsto P(\beta|k)$ captures the causal sequence $\mathcal{X} \to \mathcal{Y} \to \mathcal{Z}$.

6.2 Factorizing the Effect Disentanglement

Consider $\mathcal{Y} = \langle X, \tau \rangle \in \mathbb{R}^{b+1} \subseteq \mathbb{R}^O$ having a T-level hierarchy, with each level built up using a representation function, labeled as g_t for the t-th level. For simplicity, here, we use ω_t to represent the t-th level component of $\mathcal{Y}$ in the latent space $\mathbb{R}^L$, while its counterpart in $\mathbb{R}^{b+1}$ is denoted as Ω_t. Let the feature vector ω_t in $\mathbb{R}^L$ primarily spans a sub-dimensional space, $\mathbb{R}^{L_t}$, resulting in the spatial disentanglement sequence $\{\mathbb{R}^{L_1}, \ldots, \mathbb{R}^{L_t}, \ldots, \mathbb{R}^{L_T}\}$, which hierarchically represents $\mathcal{Y}$ with T relative timings. Function g_t maps from $\mathbb{R}^{b+1}$ to $\mathbb{R}^{L_t}$, taking into account features from all previous levels as attributes. This gives us:

$$\mathcal{Y} = \sum_{t=1}^{n} \Omega_t, \text{ where } \Omega_t = g_t\big(\omega_t;\ \Omega_1, \ldots, \Omega_{t-1}\big) \text{ with } \Omega_t \in \mathbb{R}^{b+1} \text{ and } \omega_t \in \mathbb{R}^{L_t} \subseteq \mathbb{R}^L \tag{1}$$

In the context of a purely observational hierarchy, with $\mathcal{Y}$ substituted by $Y \in \mathbb{R}^b$, The example depicted in Figure 2 (b) can be interpreted as follows: Consider three feature levels represented as $\omega_1 \in \mathbb{R}^{L_1}$, $\omega_2 \in \mathbb{R}^{L_2}$, and $\omega_3 \in \mathbb{R}^{L_3}$. For simplicity, assume each subspace is mutually exclusive, such that $L = L_1 + L_2 + L_3$. In the latent space, the triplet $\langle \omega_1, \omega_2, \omega_3 \rangle \in \mathbb{R}^L$ comprehensively depicts the image. Their observable counterparts, Ω_1, Ω_2, and Ω_3, are three distinct full-scale images, each showcasing different content. For example, Ω_1 emphasizes finger details, while the combination $\Omega_1 + \Omega_2$ reveals the entire hand.

6.3 Causal Discovery in Latent Space

Algorithm 1: Latent Space Causal Discovery

```
Result: ordered edges set E = {e_1, ..., e_n}
E = {} ; N_R = {n_0 | n_0 ∈ N, Parent(n_0) = ∅} ;
while N_R ⊂ N do
    Δ = {} ;
    for n ∈ N do
        for p ∈ Parent(n) do
            if n ∉ N_R and p ∈ N_R then
                e = (p, n);  β = {};
                for r ∈ N_R do
                    if r ∈ Parent(n) and r ≠ p then
                        β = β ∪ r
                    end
                end
                δ_e = K(β ∪ p, n) − K(β, n);
                Δ = Δ ∪ δ_e;
            end
        end
    end
    σ = argmin_e(δ_e | δ_e ∈ Δ);
    E = E ∪ σ;  N_R = N_R ∪ n_σ;
end
```

$G = (N, E)$	graph G consists of N and E
N	the set of nodes
E	the set of edges
N_R	the set of reachable nodes
$\mathbf{E}$	the list of discovered edges
$K(\beta, n)$	KLD metric of effect $\beta \to n$
β	the cause nodes
n	the effect node
δ_e	KLD Gain of candidate edge e
$\Delta = \{\delta_e\}$	the set $\{\delta_e\}$ for e
n,p,r	notations of nodes
e,σ	notations of edges

Algorithm 1 outlines the heuristic procedure for investigating edges among the initialized variable representations. We use Kullback-Leibler Divergence (KLD) as a metric to evaluate the strength of causal relationships. Specifically, as depicted in Figure 15, KLD evaluates the similarity between the relation output $\mathbf{P}(v|h)$ and the prior $\mathbf{P}(v)$. Lower KLD values indicate stronger causal relationships due to closer alignment with the ground truth. Conversely, while Mean Squared Error (MSE) is a frequently used evaluation metric, its sensitivity to data variances Reisach (2021) leads us to utilize it as a supplementary measure in this study.

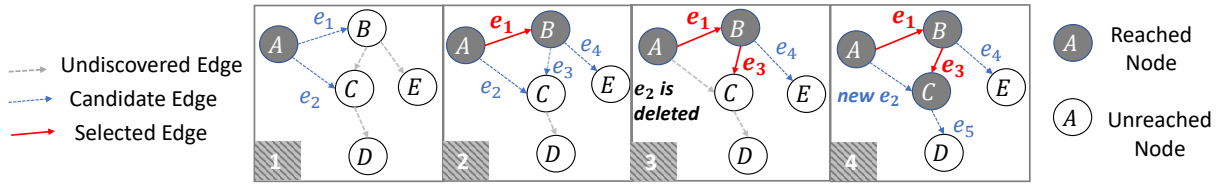


Figure 18: An example of causal discovery in the latent space.

Figure 18 illustrates the causal structure discovery process in latent space over four steps. Two edges, (e_1 and e_3), are sequentially selected, with e_1 setting node B as the starting point for e_3. In step 3, edge e_2 from A to C is deselected and reassessed due to the new edge e_3 altering C's existing causal conditions. The final DAG represents the resulting causal structure.

7 Efficacy Validation Experiments

The experiments aim to validate the efficacy of the RIRL method from three aspects: 1) the performance of the proposed higher-dimensional representations, evaluated by reconstruction accuracy, 2) the construction of a clear effect hierarchy through the stacking of relation-indexed representations, and 3) the identification of DAG structures within the latent space through discovery. A full demonstration of the conducted experiments is available online [2]. The experiments in this study present two primary limitations, detailed as follows:

Firstly, the dataset used in the current experiments may not be optimal for assessing the efficacy of RIRL. In particular, real-world causal data, such as clinical records, often contain inherent biases. While empirical constraints limited our access to such data for this study, the synthetic data we utilized may not be ideal for validating the improved model robustness conferred by RIRL. For experiments that validate the presence of such inherent biases, readers are referred to prior research Li et al. (2020).

Secondly, the time windows designated for cause and effect, l_x and l_y, are consistently set at 10 and 1, respectively. This constraint arose from an initial oversight in the experimental design, wherein the pivotal role of *dynamics* was not fully recognized, leading to restrictions set by the RNN pattern. This limitation manifests when constructing causal sequences, such as in $\mathcal{X} \to \mathcal{Y} \to \mathcal{Z}$. While the model adeptly captures single-hop effects, it struggles with two-hop information due to the dynamics in $\mathcal{Y}$ being segmented into statics by the effect window $l_y = 1$, resulting in a loss of dynamic information. However, extending the length of l_y does not pose a significant technical challenge to future works.

7.1 Hydrology Dataset

The dataset chosen for our experiments is a widely-used synthetic resource in the field of hydrology, aimed at enhancing streamflow predictions based on observed environmental conditions such as temperature and precipitation. In hydrology, deep learning, particularly RNN models, has gained favor for extracting observational representations and predicting streamflow Goodwell (2020); Kratzert (2018). We focus on a simulation of the Root River Headwater watershed in Southeast Minnesota, covering 60 consecutive virtual years with daily updates. The simulated data is from the Soil and Water Assessment Tool (SWAT), a comprehensive system grounded in physical modules, to generate dynamically significant hydrological time series.

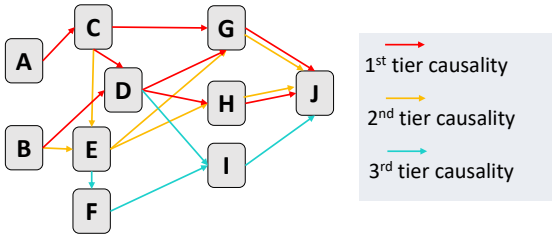
Figure 19 displays the causal DAG employed by SWAT, complete with node descriptions. The hydrological routines are color-coded based on their contribution to output streamflow. Surface runoff (1st tier) signif-

[2] https://github.com/kflijia/bijective_crossing_functions.git

Table 1: Characteristics of node attributes and their variable representation test results.

Variable	Dim	Mean	Std	Min	Max	Non-Zero Rate%	RMSE on Scaled	RMSE on Unscaled	BCE of Mask
A	5	1.8513	1.5496	-3.3557	7.6809	87.54	0.093	0.871	0.095
B	4	0.7687	1.1353	-3.3557	5.9710	64.52	0.076	0.678	1.132
C	2	1.0342	1.0025	0.0	6.2145	94.42	0.037	0.089	0.428
D	3	0.0458	0.2005	0.0	5.2434	11.40	0.015	0.679	0.445
E	2	3.1449	1.0000	0.0285	5.0916	100	0.058	3.343	0.643
F	4	0.3922	0.8962	0.0	8.6122	59.08	0.326	7.178	2.045
G	4	0.7180	1.1064	0.0	8.2551	47.87	0.045	0.81	1.327
H	4	0.7344	1.0193	0.0	7.6350	49.93	0.045	0.009	1.345
I	3	0.1432	0.6137	0.0	8.3880	21.66	0.035	0.009	1.672
J	1	0.0410	0.2000	0.0	7.8903	21.75	0.007	0.098	1.088

icantly impacts rapid streamflow peaks, followed by lateral flow (2nd tier). Baseflow dynamics (3rd tier) have a subtler influence. Our causal discovery experiments aim to reveal these underlying tiers.



ID	Variable Name	Explanation
A	Environmental set I	Wind Speed, Humidity, Temperature
B	Environmental set II	Temperature, Solar Radiation, Precipitation
C	Evapotranspiration	Evaporation and transpiration
D	Snowpack	The winter frozen water in the ice form
E	Soil Water	Soil moisture in vadose zone
F	Aquifer	Groundwater storage
G	Surface Runoff	Flowing water over the land surface
H	Lateral	Vadose zone flow
I	Baseflow	Groundwater discharge
J	Streamflow	Sensors recorded outputs

Figure 19: Hydrological causal DAG: routine tiers organized by descending causal strength.

7.2 Higher-Dimensional Variable Representation Test

In this test, we have a total of ten variables (i.e., nodes), with each requiring an individual autoencoder for initialization. Table 1 lists the statistical characteristics of their post-scaled (i.e., normalized) attributes, along with their autoencoders' reconstruction accuracies. Accuracy is assessed in the root mean square error (RMSE), where a lower RMSE indicates higher accuracy for both scaled and unscaled data.

The task is challenging due to the limited dimensionalities of the ten variables - maxing out at just 5 and the target node, J, having just one attribute. To mitigate this, we duplicate the input vector to a consistent 12-length and add 12 dummy variables for months, resulting in a 24-dimensional input. A double-wise extension amplifies this to 576 dimensions, from which a 16-dimensional representation is extracted via the autoencoder. Another issue is the presence of meaningful zero-values, such as node D (Snowpack in winter), which contributes numerous zeros in other seasons and is closely linked to node E (Soil Water). We tackle this by adding non-zero indicator variables, called *masks*, evaluated via binary cross-entropy (BCE).

Despite challenges, RMSE values ranging from 0.01 to 0.09 indicate success, except for node F (the Aquifer). Given that aquifer research is still emerging (i.e., the 3rd tier baseflow routine), it is likely that node F in this synthetic dataset may better represent noise than meaningful data.

7.3 Hierarchical Disentanglement Test

Table 2 provides the performance comparison of stacking relation-indexed representations on each node. The term "single-effect" is to describe the accuracy of a specific effect node when reconstructed from a single cause node (e.g., $B \to D$ and $C \to D$), and "full-effect" for the accuracy when all its cause nodes are stacked (e.g., $BC \to D$). To provide context, we also include baseline performance scores based on the initialized variable representations. During the relation learning process, the effect node serves two purposes: it maintains its

own accurate representation (as per optimization no.2 in 5.2) and helps reconstruct the relationship (as per optimization no.1 in 5.2). Both aspects are evaluated in Table 2.

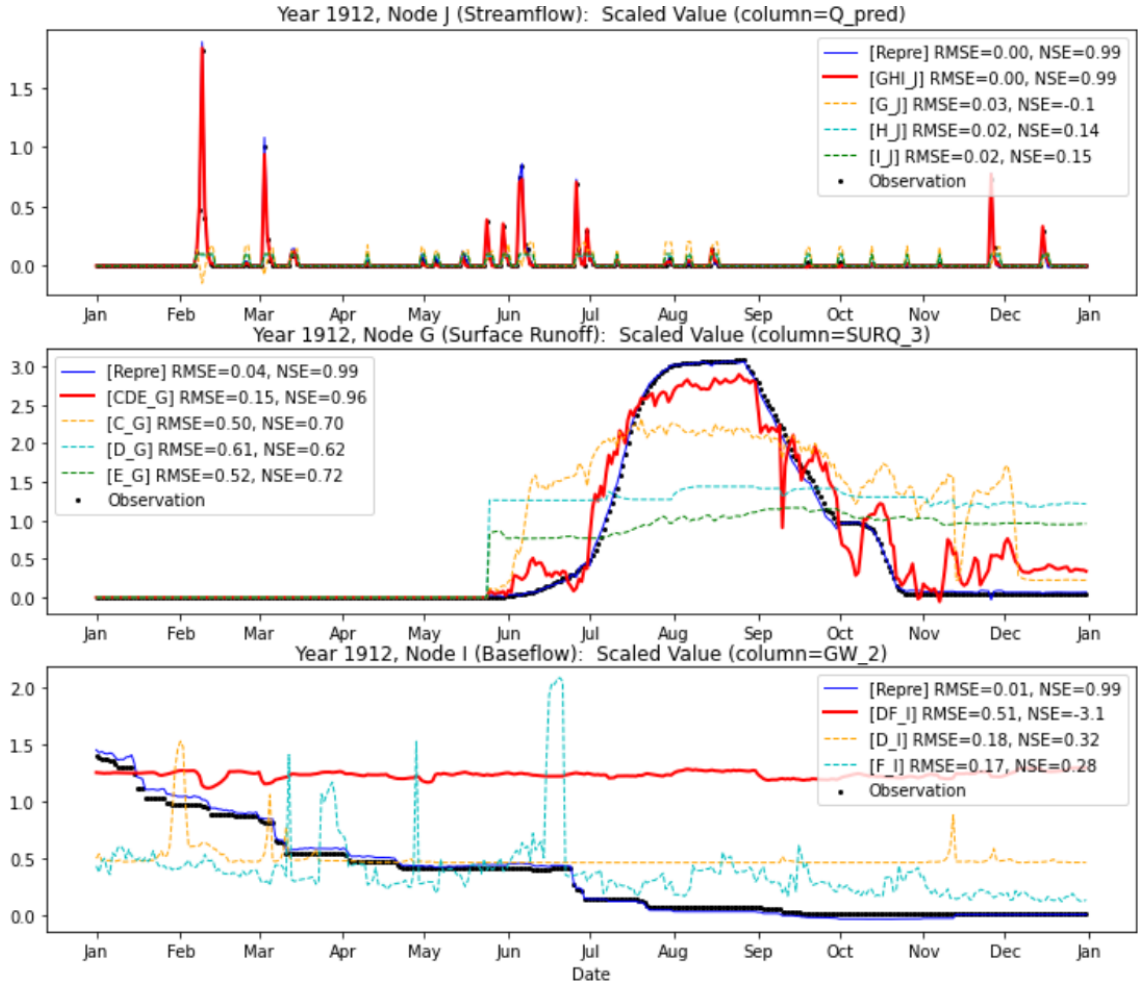


Figure 20: Reconstructed dynamical effects, via hierarchically stacked relation-indexed representations.

The KLD metrics in Table 2 indicate the strength of learned causality, with a lower value signifying stronger. For instance, node J's minimal KLD values suggest a significant effect caused by nodes G (Surface Runoff), H (Lateral), and I (Baseflow). In contrast, the high KLD values imply that predicting variable I using D and F is challenging. For nodes D, E, and J, the "full-effect" are moderate compared to their "single-effect" scores, suggesting a lack of informative associations among the cause nodes. In contrast, for nodes G and H, lower "full-effect" KLD values imply capturing meaningful associative effects through hierarchical stacking. The KLD metric also reveals the most contributive cause node to the effect node. For example, the proximity of the $C \to G$ strength to $CDE \to G$ suggests that C is the primary contributor to this causal relationship.

Figure 20 showcases reconstructed time series, for the effect nodes J, G, and I, in the same synthetic year to provide a straightforward overview of the hierarchical representation performances. Here, black dots represent the ground truth; the blue line indicates reconstruction via the initial variable representation, and the "full-effect" representation generates the red line. In addition to RMSE, we also employ the Nash–Sutcliffe model efficiency coefficient (NSE) as an accuracy metric, commonly used in hydrological predictions. The NSE ranges from -∞ to 1, with values closer to 1 indicating higher accuracy.

The initial variable representation closely aligns with the ground truth, as shown in Figure 20, attesting to the efficacy of our proposed autoencoder architecture. As expected, the "full-effect" performs better than the "single-effect" for each effect node. Node J exhibits the best prediction, whereas node I presents a challenge. For node G, causality from C proves to be significantly stronger than the other two, D and E.

Table 2: Effect Reconstruction Performances of RIRL sorted by effect nodes.

Result Node	Variable Representation (Initial) RMSE on Scaled Values	Variable Representation (Initial) RMSE on Unscaled Values	Variable Representation (Initial) BCE Mask	Cause Node	Variable Representation (in Relation Learning) RMSE on Scaled Values	Variable Representation (in Relation Learning) RMSE on Unscaled Values	Variable Representation (in Relation Learning) BCE Mask	Relationship Reconstruction RMSE on Scaled Values	Relationship Reconstruction RMSE on Unscaled Values	Relationship Reconstruction BCE Mask	Relationship Reconstruction KLD (in latent space)
C	0.037	0.089	0.428	A	0.0295	0.0616	0.4278	0.1747	0.3334	0.4278	7.6353
D	0.015	0.679	0.445	BC	0.0350	1.0179	0.1355	0.0509	1.7059	0.1285	9.6502
				B	0.0341	1.0361	0.1693	0.0516	1.7737	0.1925	8.5147
				C	0.0331	0.9818	0.3404	0.0512	1.7265	0.3667	10.149
E	0.058	3.343	0.643	BC	0.4612	26.605	0.6427	0.7827	45.149	0.6427	39.750
				B	0.6428	37.076	0.6427	0.8209	47.353	0.6427	37.072
				C	0.5212	30.065	1.2854	0.7939	45.791	1.2854	46.587
F	0.326	7.178	2.045	E	0.4334	8.3807	3.0895	0.4509	5.9553	3.0895	53.680
G	0.045	0.81	1.327	CDE	0.0538	0.9598	0.0878	0.1719	3.5736	0.1340	8.1360
				C	0.1057	1.4219	0.1078	0.2996	4.6278	0.1362	11.601
				D	0.1773	3.6083	0.1842	0.4112	8.0841	0.2228	27.879
				E	0.1949	4.7124	0.1482	0.5564	10.852	0.1877	39.133
H	0.045	0.009	1.345	DE	0.0889	0.0099	2.5980	0.3564	0.0096	2.5980	21.905
				D	0.0878	0.0104	0.0911	0.4301	0.0095	0.0911	25.198
				E	0.1162	0.0105	0.1482	0.5168	0.0097	3.8514	39.886
I	0.035	0.009	1.672	DF	0.0600	0.0103	3.4493	0.1158	0.0099	3.4493	49.033
				D	0.1212	0.0108	3.0048	0.2073	0.0108	3.0048	75.577
				F	0.0540	0.0102	3.4493	0.0948	0.0098	3.4493	45.648
J	0.007	0.098	1.088	GHI	0.0052	0.0742	0.2593	0.0090	0.1269	0.2937	5.5300
				G	0.0077	0.1085	0.4009	0.0099	0.1390	0.4375	5.2924
				H	0.0159	0.2239	0.4584	0.0393	0.5520	0.4938	15.930
				I	0.0308	0.4328	0.3818	0.0397	0.5564	0.3954	17.410

Table 3: Brief summary of the latent space causal discovery test.

Edge	A→C	B→D	C→D	C→G	D→G	G→J	D→H	H→J	B→E	E→G	E→H	C→E	E→F	F→I	I→J	D→I
KLD	7.63	8.51	10.14	11.60	27.87	5.29	25.19	15.93	37.07	39.13	39.88	46.58	53.68	45.64	17.41	75.57
Gain	7.63	8.51	1.135	11.60	2.454	5.29	25.19	0.209	37.07	-5.91	-3.29	2.677	53.68	45.64	0.028	3.384

7.4 Latent Space Causal Discovery Test

The discovery test initiates with source nodes A and B and proceeds to identify potential edges, culminating in the target node J. Candidate edges are selected based on their contributions to the overall KLD sum (less gain is better). Table 6 shows the order in which existing edges are discovered, along with the corresponding KLD sums and gains after each edge is included. Color-coding in the cells corresponds to Figure 19, indicating tiers of causal routines. The arrangement underscores the efficacy of this latent space discovery approach.

A comprehensive list of candidate edges evaluated in each discovery round is provided in Table 4 in Appendix A. For comparative purposes, we also performed a 10-fold cross-validation using the conventional FGES discovery method; those results are available in Table 5 in Appendix A.

8 Conclusions

This paper proposes a dimensionality framework to symbolize our understanding of relationships, from a novel *Relation-Oriented* perspective, and decompose our "cognitive space" accordingly, where relational knowledge is stored. It seeks to reevaluate the fundamental oversights in our current *Observation-Oriented* modeling paradigm. Specifically, based on the observational i.i.d. assumption, conventional relationship modeling intrinsically overlooks 1) the informative unobservables in $\mathbb{R}^H$, and 2) the structuralized dynamics within multi-dimensional $\mathbb{R}^T$. Instead, due to being confined within $\mathbb{R}^O$, it relies on manual specification to identify dynamical effects from observational static sequences, inherently fraught with difficulty.

Viewed through the lens of the *Relation-Oriented* framework, multifaceted issues in causality learning become unified, encompassing common confusions and concerns, from traditional causal inference to modern LLMs. Recalling the queries outlined in the Introduction, we systematically summarize our current restrictions with new insights as follows:

- ❖ *Firstly*, challenges in causal inference primarily arise from overlooking effect dynamics, due to the linear modeling constraint. This oversight leads to compensatory efforts in various aspects, such as dealing with hidden confounders and relying on the causal sufficiency assumption. Causal DAGs naturally offer a *Relation-Oriented* view; with the proposed enhancement, they can provide fundamental support.
- ❖ *Secondly*, undetectable hierarchical levels, symbolized as the hidden relation $\omega \in \mathbb{R}^H$, are inherent in our knowledge. These levels drive the need for model generalizability. With AI's capability to capture dynamics, the main challenge is incorporating structural causal knowledge to achieve generalizable causal reasoning in AI. The current paradigm struggles with identifying dynamic effects, leading to inherent biases, while transitioning to knowledge-aligned modeling suggests a shift to the new paradigm.
- ❖ *Thirdly*, although existing language models, through meta-learning, have achieved more generalizable context associations, they remain limited to observational space bound to absolute timing, and are far from structuralized "comprehension" underpinned by relative timings in our cognitive framework. Nevertheless, LLMs have demonstrated the effectiveness of meta-learning across temporal dimensional hierarchies, indicating the potential of *Relation-Oriented* meta-learning in the pursuit of AGI.

We also raise a baseline implementation of the *Relation-Oriented* paradigm, with the primary purpose of verifying the effectiveness of the "relation-indexing" methodology in extracting causal representations on demand Scholkopf (2021). Indeed, in some domains with mature structural knowledge, effective attempts have emerged under similar principles, such as the introduction of hierarchical temporal memory in neuroscience Wu (2018). The journey to achieving AGI will undoubtedly be a historically extensive and complex undertaking, necessitating a vast array of knowledge-aligned AI model constructions. This study aspires to establish foundational insights for future developments in the field.

References

Daniel L Alkon, Howard Rasmussen. A spatial-temporal model of cell activation. *Science*, 239(4843):998–1005, 1988.

Natalia Andrienko, et al. Exploratory spatio-temporal visualization: an analytical review. *Journal of Visual Languages & Computing*, 14(6):503–541, 2003.

Saurabh Arora, Prashant Doshi. A survey of inverse reinforcement learning: Challenges, methods and progress. *Artificial Intelligence*, 297:103500, 2021.

Pierre Baldi, Kurt Hornik. Neural networks and principal component analysis: Learning from examples without local minima. *Neural networks*, 2(1):53–58, 1989.

Umberto Benedetto, et al. Statistical primer: propensity score matching and its alternatives. *European Journal of Cardio-Thoracic Surgery*, 53(6):1112–1117, 2018.

Seana Coulson, et al. Understanding timelines: Conceptual metaphor and conceptual integration. *Cognitive Semiotics*, 5(1-2):198–219, 2009.

William H Crown. Real-world evidence, causal inference, and machine learning. *Value in Health*, 22(5): 587–592, 2019.

A Philip Dawid. Conditional independence in statistical theory. *Journal of the Royal Statistical Society: Series B (Methodological)*, 41(1):1–15, 1979.

Laurent Dinh, Jascha Sohl, and Samy Bengio. Density estimation using real nvp. *arXiv:1605.08803*, 2016.

Felix Elwert. Graphical causal models. *Handbook of causal analysis for social research*, pp. 245–273, 2013.

Ursula Fuller, Colin G Johnson, Tuukka Ahoniemi, Diana Cukierman, Isidoro Hernán-Losada, Jana Jackova, Essi Lahtinen, Tracy L Lewis, Donna McGee Thompson, Charles Riedesel, et al. Developing a computer science-specific learning taxonomy. *ACm SIGCSE Bulletin*, 39(4):152–170, 2007.

Dan Geiger, et al. Logical and algorithmic properties of conditional independence and graphical models. *The annals of statistics*, 21(4):2001–2021, 1993.

Clark Glymour, Kun Zhang, and Peter Spirtes. Review of causal discovery methods based on graphical models. *Frontiers in genetics*, 10:524, 2019.

Allison E Goodwell, et al. Debates—does information theory provide a new paradigm for earth science? causality, interaction, and feedback. *Water Resources Research*, 56(2):e2019WR024940, 2020.

Clive WJ Granger, et al. Modelling non-linear economic relationships. *OUP Catalogue*, 1993.

Sander Greenland, et al. Confounding and collapsibility in causal inference. *Statistical science*, 14(1):29–46, 1999.

Timothy Hospedales, Antreas Antoniou, Paul Micaelli, and Amos Storkey. Meta-learning in neural networks: A survey. *IEEE transactions on pattern analysis and machine intelligence*, 44(9):5149–5169, 2021.

Yimin Huang, Marco Valtorta. Pearl's calculus of intervention is complete. *arXiv:1206.6831*, 2012.

Aapo Hyvärinen, et al. Estimation of a structural vector autoregression model using non-gaussianity. *Journal of Machine Learning Research*, 11(5), 2010.

Saachi Jain, et al. A mechanism for producing aligned latent spaces with autoencoders. *arXiv preprint arXiv:2106.15456*, 2021.

Marcus Kaiser, et al. Unsuitability of notears for causal graph discovery. *arXiv:2104.05441*, 2021.

Frederik Kratzert, et al. Rainfall–runoff modelling using lstm networks. *Hydrology and Earth System Sciences*, 22(11):6005–6022, 2018.

Sébastien Lachapelle, Philippe Brouillard, Tristan Deleu, and Simon Lacoste-Julien. Gradient-based neural dag learning. *arXiv preprint arXiv:1906.02226*, 2019.

Brenden M Lake, et al. Human-like systematic generalization through a meta-learning neural network. *Nature*, pp. 1–7, 2023.

Jia Li, Xiaowei Jia, Haoyu Yang, Vipin Kumar, Michael Steinbach, and Gyorgy Simon. Teaching deep learning causal effects improves predictive performance. *arXiv preprint arXiv:2011.05466*, 2020.

Yunan Luo, et al. When causal inference meets deep learning. *Nature Machine Intelligence*, 2(8):426–427, 2020.

Alexander Ly, Maarten Marsman, Josine Verhagen, Raoul PPP Grasman, and Eric-Jan Wagenmakers. A tutorial on fisher information. *Journal of Mathematical Psychology*, 80:40–55, 2017.

Jianzhu Ma, et al. Using deep learning to model the hierarchical structure and function of a cell. *Nature methods*, 15(4):290–298, 2018.

Gary Marcus. The next decade in ai: four steps towards robust artificial intelligence. *arXiv preprint arXiv:2002.06177*, 2020.

Tshilidzi Marwala. *Causality, correlation and artificial intelligence for rational decision making*. World Scientific, 2015.

Mariusz Maziarz. A review of the granger-causality fallacy. *The journal of philosophical economics: Reflections on economic and social issues*, 8(2):86–105, 2015.

Allen Newell, Herbert A Simon. Computer science as empirical inquiry: Symbols and search. In *ACM Turing award lectures*, pp. 1975. 2007.

Mohammed Ombadi, et al. Evaluation of methods for causal discovery in hydrometeorological systems. *Water Resources Research*, 56(7):e2020WR027251, 2020.

Ellie Pavlick. Symbols and grounding in large language models. *Philosophical Transactions of the Royal Society A*, 381(2251):20220041, 2023.

Judea Pearl. Causal inference in statistics: An overview. 2009.

Judea Pearl. The do-calculus revisited. *arXiv preprint arXiv:1210.4852*, 2012.

Judea Pearl et al. Models, reasoning and inference. *Cambridge, UK: CambridgeUniversityPress*, 19(2), 2000.

Jonas Peters, Joris M Mooij, Dominik Janzing, and Bernhard Schölkopf. Causal discovery with continuous additive noise models. 2014.

Jonas Peters, Dominik Janzing, and Bernhard Scholkopf. *Elements of causal inference: foundations and learning algorithms*. The MIT Press, 2017.

PGMadhavan. Static dynamical machine learning – what is the difference? `https://www.datasciencecentral.com/static-dynamical-machine-learning-what-is-the-difference/`, 2016.

David Pitt. Mental Representation. In Edward N. Zalta and Uri Nodelman (eds.), *The Stanford Encyclopedia of Philosophy*. Metaphysics Research Lab, Stanford University, Fall 2022 edition, 2022.

Elad Plaut. From principal subspaces to principal components with linear autoencoders. *arXiv:1804.10253*, 2018.

Alexander G Reisach, et al. Beware of the simulated dag! varsortability in additive noise models. *arXiv preprint arXiv:2102.13647*, 2021.

Schaeffer Rylan, et al. Are emergent abilities of large language models a mirage? *arXiv preprint arXiv:2304.15004*, 2023.

Pedro Sanchez, et al. Causal machine learning for healthcare and precision medicine. *Royal Society Open Science*, 9(8):220638, 2022.

Richard Scheines. An introduction to causal inference. 1997.

Bernhard Scholkopf, et al. Toward causal representation learning. *IEEE*, 109(5):612–634, 2021.

Charles H Shea, et al. Effects of an auditory model on the learning of relative and absolute timing. *Journal of motor behavior*, 33(2):127–138, 2001.

Michael E Sobel. An introduction to causal inference. *Sociological Methods & Research*, 24(3):353–379, 1996.

Richard S Sutton, Andrew G Barto. *Reinforcement learning: An introduction*. MIT press, 2018.

Monica G Turner. Spatial and temporal analysis of landscape patterns. *Landscape ecology*, 4:21–30, 1990.

Matej Vuković, Stefan Thalmann. Causal discovery in manufacturing: A structured literature review. *Journal of Manufacturing and Materials Processing*, 6(1):10, 2022.

Yasi Wang, et al. Auto-encoder based dimensionality reduction. 184:232–242, 2016.

Naftali Weinberger and Colin Allen. Static-dynamic hybridity in dynamical models of cognition. *Philosophy of Science*, 89(2):283–301, 2022.

Gurnee Wes, Tegmark Max. Language models represent space and time, 2023.

Christopher J Wood, Robert W Spekkens. The lesson of causal discovery algorithms for quantum correlations: Causal explanations of bell-inequality violations require fine-tuning. *New Journal of Physics*, 17(3):033002, 2015.

Jia Wu, et al. Hierarchical temporal memory method for time-series-based anomaly detection. *Neurocomputing*, 273:535–546, 2018.

Gabriele Wulf, et al. Reducing knowledge of results about relative versus absolute timing: Differential effects on learning. *Journal of motor behavior*, 26(4):362–369, 1994.

Haoyan Xu, Yida Huang, Ziheng Duan, Jie Feng, and Pengyu Song. Multivariate time series forecasting based on causal inference with transfer entropy and graph neural network. *arXiv:2005.01185*, 2020.

Kun Zhang, Aapo Hyvarinen. On the identifiability of the post-nonlinear causal model. *arXiv preprint arXiv:1205.2599*, 2012.

Xun Zheng, Bryon Aragam, Pradeep K Ravikumar, and Eric P Xing. Dags with no tears: Continuous optimization for structure learning. *Advances in neural information processing systems*, 31, 2018.

Xun Zheng, Chen Dan, Bryon Aragam, Pradeep Ravikumar, and Eric Xing. Learning sparse nonparametric dags. In *International Conference on Artificial Intelligence and Statistics*, pp. 3414–3425. PMLR, 2020.

A Appendix: Complete Experimental Results of Causal Discovery

Table 4: The Complete Results of Heuristic Causal Discovery in latent space. Each row stands for a round of detection, with '#" identifying the round number, and all candidate edges are listed with their KLD gains as below. 1) Green cells: the newly detected edges. 2) Red cells: the selected edge. 3) Blue cells: the trimmed edges accordingly.

# 1	**A→ C**	A→ D	A→ E	A→ F	**B→ C**	B→ D	B→ E	B→ F								
	7.6354	19.7407	60.1876	119.7730	**8.4753**	8.5147	65.9335	132.7717								
# 2	A→ D	A→ E	A→ F	**B→ D**	B→ E	B→ F	C→ D	C→ E	C→ F	C→ G	C→ H	C→ I				
	19.7407	60.1876	119.7730	**8.5147**	65.9335	132.7717	10.1490	46.5876	111.2978	11.6012	39.2361	95.1564				
# 3	**A→ D**	A→ E	A→ F	B→ E	B→ F	**C→ D**	C→ E	C→ F	C→ G	C→ H	C→ I	D→ E	D→ F	D→ G	D→ H	D→ I
	9.7357	60.1876	119.7730	65.9335	132.7717	**1.1355**	46.5876	111.2978	11.6012	39.2361	95.1564	63.7348	123.3203	27.8798	25.1988	75.5775
# 4	A→ E	A→ F	B→ E	B→ F	C→ E	C→ F	**C→ G**	C→ H	C→ I	D→ E	D→ F	D→ G	D→ H	D→ I		
	60.1876	119.7730	65.9335	132.7717	46.5876	111.2978	**11.6012**	39.2361	95.1564	63.7348	123.3203	27.8798	25.1988	75.5775		
# 5	A→ E	A→ F	B→ E	B→ F	C→ E	C→ F	C→ H	C→ I	D→ E	D→ F	**D→ G**	D→ H	D→ I	G→ J		
	60.1876	119.7730	65.9335	132.7717	46.5876	111.2978	39.2361	95.1564	63.7348	123.3203	**2.4540**	25.1988	75.5775	5.2924		
# 6	A→ E	A→ F	B→ E	B→ F	C→ E	C→ F	C→ H	C→ I	D→ E	D→ F	D→ H	D→ I	**G→ J**			
	60.1876	119.7730	65.9335	132.7717	46.5876	111.2978	39.2361	95.1564	63.7348	123.3203	25.1988	75.5775	**5.2924**			
# 7	A→ E	A→ F	B→ E	B→ F	C→ E	C→ F	**C→ H**	C→ I	D→ E	D→ F	**D→ H**	D→ I				
	60.1876	119.7730	65.9335	132.7717	46.5876	111.2978	**39.2361**	95.1564	63.7348	123.3203	**25.1988**	75.5775				
# 8	A→ E	A→ F	B→ E	B→ F	C→ E	C→ F	C→ I	D→ E	D→ F	D→ I	**H→ J**					
	60.1876	119.7730	65.9335	132.7717	46.5876	111.2978	95.1564	63.7348	123.3203	75.5775	**0.2092**					
# 9	**A→ E**	A→ F	B→ E	B→ F	**C→ E**	C→ F	C→ I	D→ E	D→ F	D→ I						
	60.1876	119.7730	65.9335	132.7717	**46.5876**	111.2978	95.1564	63.7348	123.3203	75.5775						
# 10	A→ F	B→ E	B→ F	C→ F	C→ I	**D→ E**	D→ F	D→ I	E→ F	E→ G	E→ H	E→ I				
	119.7730	**-6.8372**	132.7717	111.2978	95.1564	**17.0407**	123.3203	75.5775	53.6806	-5.9191	-3.2931	110.2558				
# 11	A→ F	B→ F	C→ F	C→ I	D→ F	D→ I	E→ F	**E→ G**	E→ H	E→ I						
	119.7730	132.7717	111.2978	95.1564	123.3203	75.5775	53.6806	**-5.9191**	-3.2931	110.2558						
# 12	A→ F	B→ F	C→ F	C→ I	D→ F	D→ I	E→ F	**E→ H**	E→ I							
	119.7730	132.7717	111.2978	95.1564	123.3203	75.5775	53.6806	**-3.2931**	110.2558							
# 13	**A→ F**	**B→ F**	**C→ F**	C→ I	**D→ F**	D→ I	**E→ F**	E→ I								
	119.7730	**132.7717**	**111.2978**	95.1564	**123.3203**	75.5775	**53.6806**	110.2558								
# 14	C→ I	D→ I	**E→ I**	**F→ I**												
	95.1564	75.5775	**110.2558**	**45.6490**												
# 15	C→ I	D→ I	**I→ J**													
	15.0222	3.3845	**0.0284**													
# 16	**C→ I**	**D→ I**														
	15.0222	**3.3845**														

Table 5: Average performance of 10-Fold FGES (Fast Greedy Equivalence Search) causal discovery, with the prior knowledge that each node can only cause the other nodes with the same or greater depth with it. An edge means connecting two attributes from two different nodes, respectively. Thus, the number of possible edges between two nodes is the multiplication of the numbers of their attributes, i.e., the lengths of their data vectors.
(All experiments are performed with 6 different Independent-Test kernels, including chi-square-test, d-sep-test, prob-test, disc-bic-test, fisher-z-test, mvplr-test. But their results turn out to be identical.)

Cause Node	A	B		C			D			E			F	G		H	I
True Causation	A→ C	B→ D	B→ E	C→ D	C→ E	C→ G	D→ G	D→ H	D→ I	E→ F	E→ G	E→ H	F→ I	G→ J		H→ J	I→ J
Number of Edges	16	24	16	6	4	8	12	12	9	8	8	8	12	4		4	3
Probability of Missing	0.038889	0.125	0.125	0.062	0.06875	0.039286	0.069048	0.2	0.142857	0.3	0.003571	0.2	0.142857	0.0		0.072727	0.030303
Wrong Causation				C→ F				D→ E	D→ F				F→ G	G→ H	G→ I	H→ I	
Times of Wrongly Discovered				5.6				1.2	0.8				5.0	8.2	3.0	2.8	

Table 6: Brief Results of the Heuristic Causal Discovery in latent space, identical with Table 3 in the paper body, for better comparison to the traditional FGES methods results on this page.
The edges are arranged in detected order (from left to right) and their measured causal strengths in each step are shown below correspondingly. Causal strength is measured by KLD values (less is stronger). Each round of detection is pursuing the least KLD gain globally. All evaluations are in 4-Fold validation average values. Different colors represent the ground truth causality strength tiers (referred to the Figure 10 in the paper body).

Causation	A→ C	B→ D	C→ D	C→ G	D→ G	G→ J	D→ H	H→ J	C→ E	B→ E	E→ G	E→ H	E→ F	F→ I	I→ J	D→ I
KLD	7.63	8.51	10.14	11.60	27.87	5.29	25.19	15.93	46.58	65.93	39.13	39.88	53.68	45.64	17.41	75.57
Gain	7.63	8.51	1.135	11.60	2.454	5.29	25.19	0.209	46.58	-6.84	-5.91	-3.29	53.68	45.64	0.028	3.384