Probabilistic Forecasting: A Level-Set Approach

Abstract

Large-scale time series panels have become ubiquitous over the last years in areas
such as retail, operational metrics, IoT, and medical domain (to name only a few).
This has resulted in a need for forecasting techniques that effectively leverage
all available data via learning across all time series in each panel. Among the
desirable properties of forecasting techniques, being able to generate probabilistic
predictions ranks among the top. In this paper, we therefore present Level Set
Forecaster (LSF), a simple and general yet effective approach to transform a point
estimator into a probabilistic one. By recognizing the connection of our algorithm
to random forests (RFs) and quantile regression forests (QRFs), we are able to
prove consistency guarantees of our approach under certain assumptions on the
underlying point algorithm. As a byproduct, we prove the first consistency results
for QRFs under the CART-splitting criterion. Empirical experiments show that
our approach rivals state-of-the-art deep learning models in terms of forecasting
accuracy while maintaining desirable properties of tree-based algorithms. To
further support the wide applicability of our approach, we apply it to the winner of
the M5 competition for point forecasts and achieve competitive results compared
to the top 5 entries in the M5 competition for probabilistic forecasts.

1   Introduction

Tree-based methods are known to be robust, general-purpose and high-accuracy methods for general
machine learning and data science tasks, particularly those that are not image, video or text based.
This is exemplified by their popularity in competitions. In a recent interview, Anthony Goldbloom,
the CEO of Kaggle, called the prominence of gradient-boosted trees on Kaggle the "most glaring
difference" between what is used on Kaggle and what is "fashionable in academia.", a fact also
reflected in the Kaggle Data Science and Machine Learning surveys [3, 4]. In time series forecasting
in particular, tree-based methods have performed solidly in public competitions over the years [9], but
have had a recent boost in attention via the M5 accuracy competition [18], arguably the most influential
forecasting competition, which was dominated by tree-based methods. However, this recent surge in
interest is not accompanied by methodological advances. In particular, while being able to generate
probabilistic predictions ranks among the top desirable properties of forecasting techniques, the main
tree-based method that outputs predictions for multiple quantiles remains Quantile Regression Forests
(QRFs), which does not take advantage of gradient boosted trees algorithms.

In this manuscript we provide such an advancement and show its theoretical and empirical soundness:
we introduce a novel algorithm, Level Set Forecaster (LSF), which can turn any point estimator
algorithm into a probabilistic one. Applied to XGBoost [10] (in which case we refer to LSF as XLSF),
and with a simple processing component that turns time series data into tabular data, this yields one
of the first notable methodological advancement towards creating tree-based probabilistic forecasts
since the introduction of Quantile Regression Forests [21]. That said, the LSF algorithm itself is
much more general: it can take any point estimator algorithm, and apply to any tabular data.

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Our contributions can be summarized as follows:

1. We propose a novel algorithm, Level Set Forecaster (LSF), which can turn any point estimator algorithm into a probabilistic one. At a high level, it groups training data whose predictions are sufficiently close, and then uses the resulting bins of true values in the training set as the predicted distributions.

2. We prove the consistency of LSF (Theorem 1); and as a byproduct, prove the consistency of QRFs (Theorem 3). By building on the methods in [26], we introduce a general framework for consistency results that goes beyond random forests, and apply it to LSF and QRF as special cases.

3. We compare LSF with the state-of-the-art models in both tabular (see Section 6.1) and forecasting tasks (See Section 6.2), and the empirical results verify the effectiveness of the proposed approach.

The rest of the article is structured as follows. In Section 2 we introduce the Level Set Forecaster algorithm, and present the consistency of the resulting estimator in Section 3. We continue by describing the method of proof of the result in Section 4, as well as presenting the novel result (Theorem 3) for QRFs. The related work is discussed in Section 5, followed by the empirical studies for tabular and time series data in Section 6.

2 Level Set Forecaster

Given a dataset $\mathcal{D} := \{(x_i, y_i)\}_{i=1,...,n} \subset \mathbb{R}^d \times \mathbb{R}$, we train a given point prediction algorithm $\mathcal{A}$ on $\mathcal{D}$ to arrive at a model $f : \mathbb{R}^d \to \mathbb{R}$. Our goal is to output a probabilistic predictor based on f, such that one could query an arbitrary quantile at any testing point $x \in \mathbb{R}^d$. In more precise terms, assuming (x_i, y_i)'s are sampled iid from the joint distribution (X, Y), we wish to find an estimator for $P(Y \leq y | X)$.

In Algorithm 1, we propose a conceptually simple yet effective approach to turn f into a probabilistic forecaster. At a high level, Algorithm 1 creates groupings of training data points such that their predictions are "sufficiently close." These groupings are used to partition the feature space. With a new testing data point x, we find the partition cell whose predictions are "close to" $f(x)$, and the empirical samples that belong to the same partition cell yield the predictive distribution.

Forming the partition of the feature space. Algorithm 1 has a natural geometric interpretation. First, one trains the algorithm $\mathcal{A}$ on the data $\mathcal{D}$ to get a model $f : \mathbb{R}^d \to \mathbb{R}$. The level sets $\{f^{-1}(f(x_i))\}$ are disjoint, but they do not in general satisfy that their union covers $\mathbb{R}^d$, because it may be that $\{f(x_1), ..., f(x_n)\}$ is not equal to the entire image of f. We therefore create Voronoi sets for $\{f(x_1), ..., f(x_n)\}$, and look at inverse images of these. Namely, writing $\{f(x_1), ..., f(x_n)\}$ as $\{v_1, ..., v_k\}$ with $v_1 < \cdots < v_k$, the jth Voronoi set is:

$$V_j = \{a \in \mathbb{R} | v_j = \mathrm{argmin}_{\{v_1,...,v_k\}}(|v_i - a|)\}. \tag{1}$$

We now define $\tilde{B}_j = f^{-1}(V_j)$, with each $\tilde{B}_j$ being a disjoint union of level sets of f. The inverse Voronoi sets $\tilde{B}_j$'s are disjoint, and their union covers all of $\mathbb{R}^d$. This is not yet, however, our final partition: with the basic set-up described above, the cardinality of each bin $|\{y_i \in \{y_1, ..., y_n\} | x_i \in \tilde{B}_j\}|$ may be too small (often equal to 1) to ensure the desired property of consistency. In the final part of the algorithm we merge the different $\tilde{B}_j$ together, by going sequentially over $\tilde{B}_1, ..., \tilde{B}_k$ (associated with the ascending values $v_1, ..., v_k$), grouping them until a sufficient bin size is ensured, and then proceeding to a new grouping, arriving at the final partition $\mathbb{R}^d = \bigcup B_j$, where each B_j is a union of $\tilde{B}_i$'s. The intuition behind the partitioning algorithm is that for a test sample (X, Y) and a value $y \in \mathbb{R}$ we expect $P(Y \leq y | X)$ to not vary wildly as X runs over a particular partition cell B_j.

Using the partition to make inferences. In the notation of Algorithm 1, the algorithm finds the closest element v_j in $[v_1, ..., v_k]$ to $f(x)$ (by binary search), and then computes and caches the quantile of `pred_to_bin`$[v_j]$.

Algorithm 1: Level Set Partitioning Algorithm

Input: $\mathcal{D} := \{(x_i, y_i)\}_{i=1,...,n} \subset \mathbb{R}^d \times \mathbb{R}$, a natural number `min_bin_size`, a point prediction algorithm $\mathcal{A}$ (e.g., XGBoost).

Train $\mathcal{A}$ on $\mathcal{D}$, and call the resulting model $f : \mathbb{R}^d \to \mathbb{R}$.

Create a dictionary `pred_to_bin` whose keys are $\{f(x_i)\}_i$, such that the value of $f(x_i)$ is the list of y_j satisfying $f(x_j) = f(x_i)$.

Sort the distinct elements of $\{f(x_1), ..., f(x_n)\}$ in ascending order, and denote them $[v_1, ..., v_k]$.

Initialize an empty dictionary `res_pred_to_bin`, and an empty list `current_bin`.

 # If the size of a bin is smaller than the minimum required size, successively merge it with the bins on its right (higher value) until the minimum size constraint is satisfied.

for v in $[v_1, ..., v_k]$ **do**

    Concatenate `pred_to_bin[v]` at the end of `current_bin`.

    `res_pred_to_bin[v]` $\leftarrow$ `current_bin` (pass `current_bin` by reference)

    if `len(current_bin)` $\geq$ `min_bin_size` **then**

        Point `current_bin` to a new empty list object.

 # If the last bin is too small, merge to the left.

if `len(res_pred_to_bin[`v_k`])` $<$ `min_bin_size` **then**

    concatenate `res_pred_to_bin[`v_i`]` to `res_pred_to_bin[`v_k`]`, where i is the maximal index satisfying that `res_pred_to_bin[`v_i`]` $\neq$ `res_pred_to_bin[`v_k`]`.

`pred_to_bin` $\leftarrow$ `res_pred_to_bin`

Output: A partition such that $w_1, w_2 \in \mathbb{R}^d$ are said to be in the same partition if `pred_to_bin` maps $\mathrm{argmin}_{\{v_1,...,v_k\}}(|v_i - f(w_1)|)$ and $\mathrm{argmin}_{\{v_1,...,v_k\}}(|v_i - f(w_2)|)$ to the same bin.

In more mathematical terms, Algorithm 1 outputs a partition $\mathbb{R}^d = \bigsqcup B_j$ of the feature space into finitely many disjoint sets. Queried at a feature vector $x \in B_l \subset \mathbb{R}^d$, and for any value $y \in \mathbb{R}$, our probabilistic forecasting algorithm returns an estimator for $P(Y \leq y|X)$ at $X = x$:

$$\hat{\eta}_{LSF}(x) := \frac{|\{\gamma \in L_l | \gamma \leq y\}|}{|L_l|}, \tag{2}$$

where $L_l := \{y_i \in \{y_1, ..., y_n\} | x_i \in B_l\}$ is the associated true target values for the partition cell B_l. We shall study the consistency property of the estimator $\hat{\eta}_{LSF}$ in the next section.

Remarks. Algorithm 1 resembles the classic Quantile Regression Forecasts (QRFs) ([21]) in the following sense. In QRFs, the random forests are trained in the regular sense, i.e., one splits the nodes according to the CART-splitting criterion. The probabilistic predictions are generated by "opening" the leaf nodes of each tree, and empirically sampling the predictions instead of predicting with the average values in the leaves. Here, the second step is exactly the same while the "leaf nodes" are replaced by partitions induced by the level sets of an arbitrary point estimator. In fact, if one used a decision tree for $\mathcal{A}$ and let `min_bin_size` be 1 then Algorithm 1 reduces to a QRF with one tree.

We refer to the estimator described via Equation 2 with the partitioning algorithm described as in Algorithm 1 as the Level Set Forecaster (LSF) associated to $\mathcal{A}$. In the case that $\mathcal{A}$ is XGBoost, we refer to it as XLSF. We remark that $\mathcal{A}$ can theoretically be chosen to be continuous as well, and in that case collisions $f(x_i) = f(x_j)$ will likely never occur, and occur in practice only due to the machine's precision limitations.

3 Main Theorem: LSF is Consistent

In this section, we aim to provide a quantitative understanding of the Level-set Forecaster (LSF), in particular, the consistency of LSF. The key insight that motivated the LSF algorithm was the observation that the literature regarding the consistency of Random Forests (RFs) can be extended to apply in much more general contexts.

We remark that it is precisely because the partitions of trees grown by the CART-splitting algorithm are data-dependent that consistency of RFs has only been proven recently [26], while most other work

([7, 28, 6, 21, 25]) has focused on easier algorithms that are not used in practice. We can leverage these methods to prove consistency of LSF as well (Theorem 1); and as a byproduct we also prove consistency of QRFs under the CART-splitting algorithm (Theorem 3).

In order to have a consistency assurance, we require mild assumptions on the data generating process, and some intuitive assumptions on the point forecasting algorithm $\mathcal{A}$.

Assumption 1.

1. **Data generating process assumption**: *There exist pdfs of $P(Y \leq y|X)$ as y varies that form a uniformly equicontinuous family of functions.*

2. **Assumptions on the point estimator** $\mathcal{A}$: *For each n let f_n be the model resulting from training $\mathcal{A}$ on $\mathcal{D}$, and let $W_n := \{f_n(X_1), ..., f_n(X_n)\}$. Assume that $\mathcal{A}$ has a choice (for every n) of hyperparameters that would satisfy:*

   (a) *Collisions among the training data $f_n(X_i) = f_n(X_j)$ are rare in the precise sense that there exists a positive number C such that:*

   $$P\left(C < \frac{|W_n|}{n}\right) \to 1.$$

   (b) *For X an independent variable following the distribution of the X_i's the value $f_n(X)$ does not tend to be extremal among the $f_n(X_i)$'s. To be precise, for any sequence $d_n \to 0$ of positive numbers we have that*

   $$P\left(\frac{|\{v \in W_n|f_n(X) < v\}|}{|W_n|} > d_n\right) \to 1, P\left(\frac{|\{v \in W_n|v < f_n(X)\}|}{|W_n|} > d_n\right) \to 1.$$

   (c) *The image of f_n is "dense in probability". To be precise,*

   $$\max_{i=2,...,t_{n-1}} \text{Vol}(V_i) \to 0$$

   in probability, where the V_i's are the Voronoi sets from Equation 1.

   (d) *Inverse images of small volume sets via f_n have a small volume in probability: If $P_n \subset \mathbb{R}^n$ are random closed subsets satisfying that $\text{Vol}(P_n) \to 0$ in probability, then $\text{Vol}(f_n^{-1}(P_n)) \to 0$ in probability.*

We are able to achieve the following consistency guarantee for LSF:

Theorem 1. *Under Assumption 1, letting* `min_bin_size`$= (\ln(n))^2$, *LSF is mean square consistent. That is, for any value $y \in \mathbb{R}$, we have*

$$\mathbb{E}(|\hat{\eta}_{LSF}(X) - P(Y \leq y|X)|^2) \to 0, \tag{3}$$

where the convergence is uniform in y.

As a byproduct of proving this result, we also give the first consistency result for QRFs grown under the CART-splitting algorithm under an additive regression model assumption. As this involves more notation, we postpone the statement of this result to Theorem 3.

Remark 1. Acute readers would immediately recognize the shortcomings of Theorem 1, i.e., we assume that $\mathcal{A}$, with its chosen hyperparameters, satisfies the intuitive assumptions in its statement. However, we have yet to prove that for a particular choice of $\mathcal{A}$ (for example, XGBoost) the properties hold. We expect the assumptions would hold for a reasonable choice of $\mathcal{A}$, but the precise statement is left as a future work.

4 Method of Proof: Extend Results on RF Consistency to More General Settings

Our strategy for proving Theorem 1 is to generalize and strengthen existing results on the consistency of RFs, in particular through the work of [26]. Next, we propose a unified framework that subsumes both the standard RFs (and QRFs) and the proposed LSF estimator, allowing us to arrive at the consistency results. We begin by introducing additional notation and nomenclature.

4.1 Problem Setup and Existing Results

Define a *data-based partitioning algorithm $\mathcal{B}$* as being any algorithm that takes univariate training data (a finite subset $\mathcal{D} = \{(X_1, Y_1), ..., (X_n, Y_n)\}$ of $\mathbb{R}^d \times \mathbb{R}$) and outputs a (potentially random) finite set of disjoint subsets B_j of $\mathbb{R}^d$ whose disjoint union is $\mathbb{R}^d$. Examples include Algorithm 1, as well as any tree-growing algorithm such as the CART-splitting algorithm (see Algorithm 2, recalled below). (Note that Algorithm 1 does not inherently have an underlying tree structure, even if $\mathcal{A}$ is chosen to be a tree-based forecaster such as XGBoost.)

We define a mean-regression (resp. quantile-regression) $\mathcal{B}$-estimator, trained on $\mathcal{D}$ as an estimator of $\mathbb{E}(\xi(Y)|X)$, where ξ is the identity map $\xi(Y) := Y$ (resp. $\xi(Y) := I_{Y \leq y}$) given by first applying $\mathcal{B}$ to get a partition $\mathbb{R}^d = \bigcup B_j$; and then at inference taking a feature vector $x \in B_l$ to: $\frac{\sum_{a \in L_l} \xi(a)}{|L_l|}$, where $L_l := \{Y_i \in \{Y_1, ..., Y_n\} | X_i \in B_l\}$. Now we are ready to introduce the central object of our study, *generalized RF/QRF*,

Definition 1. *A Generalized RF (resp. Generalized QRF) grown using the partitioning algorithm $\mathcal{B}$ with M estimators is defined as:*

$$\hat{\eta}(x) := \frac{\sum_{j=1}^{M} \hat{\eta}_j(x)}{M}, \tag{4}$$

where $\hat{\eta}_j$ is a mean-regression (resp. quantile-regression) $\mathcal{B}$-estimator trained on only part of the data: for the jth estimator subsample uniformly without replacement a_n many datapoints from $\mathcal{D}$ to train on, where a_n is a (non-random) sequence of natural numbers satisfying $1 \leq a_n \leq n$.

We remark that LSF is simply a generalized QRF grown using Algorithm 1 with one estimator and full subsampling; and that RFs (resp. QRFs) are simply generalized RFs whose partitioning algorithm is induced by a tree-growing algorithm such as the CART splitting algorithm (see Algorithm 2).

To fix some notation, we let t_n (a random variable that can depend on the data) be the number of partitions in the first $\mathcal{B}$-estimator, and we let $R_{x,n}$ be the partition in the first $\mathcal{B}$-estimator that contains a feature vector $x \in \mathbb{R}^d$.

The main question we consider is:

Problem 1. *Under what assumptions is the generalized RF/QRF estimator $\hat{\eta}(\cdot)$ constructed in (4) L^2 (mean square) consistent? i.e.,*

$$\mathbb{E}(|\hat{\eta}(X) - \mathbb{E}(\xi(Y)|X)|^2) \to 0,$$

where ξ is defined either as $\xi(T) = T$ or $\xi(T) = I_{T \leq y}$.

Our investigation starts with the recent work on the consistency of random forests. In Theorem 1 of [26] the authors proved mean square consistency in the mean regression case ($\xi(T) = T$) under some mild assumptions, where the partitioning algorithm is the one induced by a tree grown via the CART splitting criterion.

We briefly recall the tree-growing algorithm for the CART splitting criterion:

Algorithm 2: Mean and Quantile Regression Forests with the CART-Splitting Criterion; CART loss is recalled in Appendix A. Note that unlike Algorithm 1 the number of partition cells t_n here is a hyper-parameter and therefore not random.

Input: We fix a natural number $m_{try} \leq d$; and sequences of natural numbers $\{t_n\}_n$, $\{a_n\}_n$ such that $t_n \leq a_n \leq n$.;
while *not reaching t_n many leaves* **do**
> For the subsampling of each tree, choose a_n points uniformly without replacement from the training set.
> At each step, and for each leaf, choose the split that minimizes CART loss, considering only m_{try} split dimensions sampled uniformly without replacement from $\{1, ..., d\}$.

Mean prediction: output the average value of the leaf nodes. ($\xi(T) = T$);
Quantile prediction: output the average value of the leaf nodes. ($\xi(T) = I_{T \leq y}$)

In [26], the authors make the following assumptions:

Assumption 2 (Assumptions in [26])**.**

1. *Subsamples and tree growth are same as in the CART-splitting algorithm.*

2. *Additive Regression Model Assumption: Y is normal with fixed variance σ^2 and with mean $\sum_{i=1}^{d} m_i(p_i(X))$, where p_i is the projection onto the ith coordinate, where each $m_j : [0,1] \to \mathbb{R}$ is a continuous function, and where the marginal distribution of X is uniform in the hypercube $[0,1]^d$.*

3. *Assume that $\lim_{n\to\infty} \dfrac{t_n \ln(a_n)^9}{a_n} = 0, \lim_{n\to\infty} t_n = \infty$.*

The result of primary interest in [26] is the following.

Theorem 2. *([26], Theorem 1) Under Assumptions 2 we have mean square consistency*

$$\mathbb{E}\left[|\hat{\eta}_1(X) - \mathbb{E}(Y|X)|^2\right] \to 0,$$

where $\xi(T) := T$ is the identity transformation.

Remark 2. [26] phrase their result as the consistency of $\mathbb{E}(\hat{\eta}_1(X)|X, \{X_i, Y_i\}_i)$, i.e. the case of an infinite number of trees, but they in fact reduce first to the stronger statement that a single tree is consistent.

In [21], the quantile regression case ($\xi(T) = I_{T \leq y}$) was addressed, but only for the case of label-independent weights, which precludes the case of the partitioning algorithm being induced by trees grown via the CART-splitting algorithm.

4.2 Consistency of Generalized RFs/QRFs

To begin with, we state a novel result, the first consistency proof for QRFs under the CART-splitting criterion.

Theorem 3. *Under Assumption 2, with the assumption 2.3 replaced by the weaker assumption that $\lim_{n\to\infty} \dfrac{t_n \ln(a_n)}{a_n} = 0$, $\lim_{n\to\infty} t_n = \infty$, we have that:*

$$\mathbb{E}(|\hat{\eta}_1(X) - P(Y \leq y|X)|^2) \to 0,$$

uniformly in y, where $\xi(T) := I_{T \leq y}$.

We remark that convergence in probability (the notion used in [21]) and L^2 convergence coincide for QRFs because the estimator and the CDF are both bounded.

Both Theorems 3 and Theorem 1 are corollaries from the following vast generalization of Theorem 2:

Theorem 4. *A generalized QRF with a single estimator (a "generalized quantile regression tree") is mean square consistent under the following conditions:*

1. *The data generating process follows Assumption 1.1.*

2. *There exists a (potentially random) assignment of a hyperrectangle $D_x \subset \mathbb{R}^d$ containing x for each $x \in \mathbb{R}^d$, such that the probability that $\mathbb{E}(Y|X)$ is almost surely constant over D_x is 1, and such that $\mathrm{Vol}(R_{X,n} \backslash D_X)$ approaches 0 in probability.*

3. *(Only needed if one does not choose $D_x = \{x\}$.) If $\mathbb{E}(Y|X)$ is almost surely constant over a hyperrectangle, then $P(Y \leq y|X)$ is almost surely constant over this hyperrectangle for all $y \in \mathbb{R}$.*

4. *$\dfrac{t_n \ln(a_n)}{a_n} \to 0$ in probability.*

Namely, under these conditions, we have that:

$$\mathbb{E}\left[|\hat{\eta}_1(X) - P(Y \leq y|X)|^2\right] \to 0,$$

where the convergence is uniform in y, and where $\xi(T) := I_{T \leq y}$.

220 Theorem 1 follows from Theorem 4 as an immediate corollary:

221 *Proof.* By Assumption 2a, and our choice `min_bin_size`$= (\ln(n))^2$, it follows that in probability
222 t_n has the same order of magnitude as $\frac{n}{(\ln(n))^2}$. It is therefore clear that Assumption 4 in Theorem 4
223 is satisfied.

224 Note that $R_{X,n}$ is the union of at most $(\ln(n))^2$ many $f^{-1}(V_j)$. By Assumption 2b, with probability
225 converging to 1 none of these V_j's are V_1 nor V_{t_n}. Now Assumptions 2c and 2d immediately imply
226 that $\text{Vol}(R_{X,n})$ goes to 0 in probability. This implies Assumption 2 in Theorem 4 for the assignment
227 $D_x := \{x\}$. □

228 It is less obvious to see why Theorem 4 implies Theorem 3. The assertion boils down to proving that
229 the assumptions of Theorem 3 imply Assumption 2 of Theorem 4. The assignment $x \mapsto D_x$ in this
230 situation requires the notion of theoretical cuts, and we refer to Appendix B for its precise definition
231 and for the proof of Theorems 3 and 4.

5   Related Work

233 The classical conformal prediction algorithm [27], whose underlying principle is bootstrapping
234 the residuals on a validation set, is another approach for turning point estimator algorithms into
235 probabilistic ones. The natural setting in this approach is to create prediction intervals, rather than
236 to make predictions for particular quantiles. Unlike LSF, the prediction intervals from conformal
237 predictions are always symmetric about the point prediction with a fixed length. More sophisticated
238 conformal prediction algorithms exist that allow for variable length prediction intervals (e.g. [23],
239 which takes as input two quantile regression estimators), but they no longer satisfy that they take a
240 single point estimator algorithm to a probabilistic one. We refer to Appendix C for more details.

241 The main available probabilistic tree-based algorithm other than LSF (applied to a tree-based point
242 estimator algorithm) that does not require a list of quantiles during training is QRFs [21]. We remark
243 that LSF applied to a decision tree is equivalent to a QRF with a single tree, and that therefore QRFs
244 with multiple trees are an ensemble of these simple LSF algorithms.

245 Another approach for creating tree-based probabilistic predictions is to learn the parameters of a
246 family of conditional distributions (e.g., [8]). For tree-based methods, this has been proposed [19, 20].
247 We do not further compare against this method in our experiments because the implementations are
248 not available and instead focus on comparisons with state-of-the-art methods for forecasting that use
249 a similar approach (e.g., DeepAR [24]).

250 For probabilistic time series forecasting, a number of global [14] models are available (e.g., [5, 22,
251 11, 17]). In contrast to classical forecasting models (e.g., [13]) which are local (parameters are
252 estimated per time series independently), tree-based methods are best employed as global models
253 (learning parameters over the entire panel of time series). In our empirical comparisons, we chose
254 DeepAR [24] and CNN-QR [29] as global, deep-learning based models for their popularity and
255 robust general-purpose performance.

6   Experiments

257 In Section 6.1 we compare XLSF (LSF with XGBoost) against QRFs (as a leading tree-based
258 probabilistic algorithm) and Conformalized Predictions (as a baseline for turning point estimators
259 into probabilistic ones). In Section 6.2 we apply LSF to time series data. All experiments are done
260 using Amazon Sagemaker [16] with instance type `ml.m4.16xlarge`. All hyperparameters used are
261 specified in Appendix F.

6.1   Experiments on Tabular Data

263 A main competitor for LSF for using tree-based methods to create probabilistic predictions is QRFs
264 ([21]). In the experiments we use the skgarden implementation ([2]). QRFs are made up of multiple
265 trees, but XLSF is one generalized tree made up of an XGBoost ensemble. It is therefore interesting
266 to see how the two compare. By design, conformal predictions are phrased in terms of prediction

interval (a pair of quantiles) rather than a range of quantiles. For example for an $\alpha = 0.1$, conformal predictions aim to estimate P05 and P95 so as to have an accuracy of the prediction intervals (percent of times that the true value is in the prediction interval) that is at least 90%. We therefore chose only the pair of quantiles P05 and P95 in our experiments, as opposed to 3 quantiles as in Section 6.2.

The results are shown in Table 1, and we see that XLSF outperforms conformal predictions by a wide margin on quantile prediction, though conformal predictions' prediction interval accuracy was more faithfully close to 90%. We remark that to reap the benefits of both algorithms, one can feed XLSF into the "Conformalized Quantile Prediction" algorithm; see [23], and the discussion in Appendix C. We also remark that QRF is much slower compared to XLSF, as shown in Table 1. The datasets in this section were small, but in Section 6.2 we will see that this makes classical QRF impractical for time series predictions, while XLSF is quite competitive.

	XLSF				QRF				Conformalized Predictions			
	P05	P95	accuracy	time (s)	P05	P95	accuracy	time (s)	P05	P95	accuracy	time (s)
facebook$_1$	0.103	**0.288**	94.31%	7.668	**0.094**	0.317	92.74%	29.918	0.471	0.428	89.98%	6.942
facebook$_2$	0.097	**0.293**	95.32%	14.790	**0.094**	0.299	92.93%	103.61	0.391	0.445	89.89%	14.442
meps$_{19}$	**0.100**	**0.562**	93.37%	5.874	0.111	0.687	89.45%	9.159	0.633	0.766	89.64%	5.305
meps$_{20}$	**0.100**	**0.664**	92.78%	6.384	0.107	0.674	88.37%	10.526	0.498	0.774	90.13%	5.884
meps$_{21}$	**0.100**	**0.555**	92.68%	6.215	0.105	0.635	88.40%	9.111	0.525	0.802	89.94%	5.246
concrete	0.036	0.039	76.69%	0.627	0.036	**0.037**	82.52%	0.167	**0.031**	0.042	86.89%	0.089
star	**0.011**	0.012	78.52%	1.004	**0.011**	0.014	79.21%	0.500	**0.011**	**0.011**	87.99%	0.305
bio	0.082	0.132	87.63%	13.195	**0.073**	**0.096**	84.40%	33.153	0.103	0.128	89.99%	4.528
community	0.105	0.189	76.19%	1.120	**0.079**	0.187	86.71%	1.132	0.136	**0.184**	92.48%	0.955
bike	0.056	0.059	87.74%	3.403	**0.043**	**0.044**	80.53%	2.828	0.050	0.048	91.23%	0.734

Table 1: Benchmarking results: the datasets were taken from `https://github.com/yromano/cqr/tree/master/datasets`. Accuracy is percent of times the true value was in the prediction interval. (It should revolve around 90%.) P05 and P95 are weighted quantile losses.

6.2 LSF for Time Series Prediction

In order to apply LSF to time series data, a few choices need to be made. First, the method for preprocessing the data into tabular data; second, whether to use a single model for the entire forecast horizon, or one model for each timestep; and finally, one has to decide on what point estimator algorithm $\mathcal{A}$ to use and its hyperparameters. The winning solution to the M5 accuracy competition has already made all of these decisions, and so we can simply wrap it with LSF, and see how it compares with the top 5 solutions for the M5 uncertainty competition. We also propose a simple default choice (using only lag features, one model for each timestep, and using XGBoost) and apply it to multiple datasets.

M5 competition The M-competitions are a series of competitions led by Spyros Makridakis intended to evaluate and compare the accuracy of different forecasting methods. The most recent competition, the M5 competition [18], was in fact a twin competitions: an accuracy competition (point forecasting) and an uncertainty competition (probabilistic forecasting) with the same training data. The winning solution (by YeonJun-IN, available in [1]) to the M5 accuracy competition is a tree-based solution that trains 220 lightgbm ([15]) models (with different feature engineering set ups and/or trained on different parts of the data) and combines them to create a point forecast. The preprocessing involved more than mere lag features.

We remark that the M5 uncertainty competition requires forecasts not only of the individual time series, but also of hierarchical aggregates. In our reported metrics we only consider the bottom level of the hierarchy, since there is no trivial way to produce aggregate forecasts using the LSF-wrapping of the winning M5 accuracy competition's solution.

We can see that wrapping the M5 accuracy competition's winning solution is competitive among the top 5 solutions for the M5 uncertainty competition, showing the out-of-the-box we were able to turn quality point-forecasts into quality probabilistic forecasts.

quantile	LSF-wrapping of winning M5 solution	Winner	Runner-up	Third	Fourth	Fifth
0.005	0.01227	0.010	0.04201	0.01891	0.01225	0.01133
0.025	0.05484	0.05004	0.08649	0.06446	0.05483	0.05739
0.165	0.31677	0.31276	0.33782	0.33371	0.31566	0.35329
0.25	0.45108	0.44436	0.46158	0.46699	0.44937	0.49415
0.5	0.72434	0.69886	0.69044	0.74712	0.71279	0.7661
0.75	0.72848	0.68747	0.72231	0.69522	0.71071	0.70321
0.835	0.60748	0.58519	0.59897	0.59059	0.61292	0.59058
0.975	0.21801	0.18236	0.19664	0.19268	0.19348	0.20365
0.995	0.0863	0.0551	0.07689	0.07433	0.06225	0.06811
mean_wQL	0.35551	**0.33624**	0.35702	0.35378	0.34714	0.36087

Table 2: The LSF-wrapping of the winning M5 accuracy competition's solution on the top point forecast submission compared with top results in the uncertainty competition. Note that the non-monotonous nature of the metrics for the winning solutions is due to the fact that it is being evaluated only on the base level, and that the metric used is different from the competitions' metric.

Benchmarking datasets. We now turn to some benchmarking with certain default choices of preprocessing and inference on common time series datasets. To be precise, we preprocess the time series data into tabular data by sampling P context windows with only lag features, and then training a separate probabilistic forecaster for each time-step in the future. (See Algorithm 3 in the appendix for more details; see also Appendix F.)

We refer to this set-up as Tree-based Time Series Wrapper (TTSW). We remark that by design, TTSW with XLSF never predicts above (resp. below) the maximal (resp. minimal) value observed. We also include TTSW with QRFs instead of XLSF, and with lightgbm ([15]) with quantile loss, which we refer to as Quantile Regression. If one chooses XLSF or QRF there is no need to specify quantiles at training time, but for Quantile Regression, one specifies quantiles before training, and for each timestep in the forecast horizon it trains as many models as the number of quantiles requested.

	TTSW (XLSF)	TTSW (QR)	TTSW (QRF)	DeepAR	CNN-QR
electricity	**0.0499**	0.1478	0.0785	0.0592	0.0601
parts	0.9709	0.7584	0.9617	0.9877	**0.7566**
m4_daily	**0.0148**	0.0281	0.0188	0.0239	0.0154
traffic	0.1212	0.1062	0.1289	**0.0913**	0.1134
wiki10k	0.2926	0.2534	0.3069	**0.2421**	0.2482
dcrideshare	0.3381	0.3171	0.3929	0.2837	**0.2820**

Table 3: Mean Weighted Quantile Loss (across P10, P50, P90)

By default $P = 1000000$ and the context window is equal to the forecast horizon. In Table 3, since QRF is much slower to train, we had to reduce P to 10000, which means it had 100 times less data-points to train on. Even with this adjustment, it took considerably longer than the other methods. Our results in Table 3 show that TTSW (both XLSF and with Quantile Regression) perform competitively with state-of-the-art deep learning forecasting methods ([24, 29]), while maintaining advantages of tree-based methods such as interpretability. For detailed results, please refer to Appendix E.

7 Broader Impact

The present article stems from the authors' work on time series forecasting in industrial settings. The proposed methods are applicable to forecasting and allow to re-use multiple time series panels. More accurate forecasts allow for better decision making which we hope will lead to benefits such as waste reduction, reduction of emission through improve transportation costs and optimisation of energy consumption.

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