Greedier is Better: Selecting Multiple Neighbors per Iteration for Sparse Subspace Clustering

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Abstract

Sparse subspace clustering using greedy-based neighbor selection, such as orthogonal matching pursuit (OMP), has been known as a popular computationally-efficient alternative to the standard ℓ_1-minimization based methods. This paper proposes a new SSC scheme using generalized OMP (GOMP), a soup-up of OMP whereby multiple, say $p(\geq 1)$, neighbors are identified in each iteration, along with a new stopping rule requiring nothing more than a knowledge of the ambient signal dimension and the number p of neighbors identified per iteration. Compared to conventional OMP (i.e., $p = 1$), the proposed GOMP method involves fewer iterations, thereby enjoying lower algorithmic complexity; in addition, the proposed stopping rule is free from an off-line estimation of the subspace dimension or noise strength. Under the semi-random model, analytic performance guarantees are provided. It is shown that, with a high probability, (i) GOMP can retrieve more true neighbors than OMP, consequently yielding higher data clustering accuracy, and (ii) the proposed stopping rule halts neighbor search once the number of recovered neighbors is close to the subspace dimension. Issues about selecting p for practical implementation are also discussed. Computer simulations using both synthetic and real data are provided to demonstrate the effectiveness of the proposed approach and validate our analytic study.

1 Introduction

1.1 Motivation

Subspace clustering (Vidal, 2011; Yang et al., 2008; Goh & Vidal, 2007) is a key enabling technique in modern unsupervised machine learning and its principles can be recapitulated as follows. Consider a noisy dataset $\mathcal{Y} = \{\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N\} \subset \mathbb{R}^n$ whose ground truth obeys a disjoint union as

$$\mathcal{Y} = \mathcal{Y}_1 \cup \mathcal{Y}_2, ... \cup \mathcal{Y}_L, \tag{1}$$

where cluster $\mathcal{Y}_k \subset \mathbb{R}^n$ consists of $|\mathcal{Y}_k| > 0$ noisy data points coming from a d_k-dimensional subspace $\mathcal{S}_k$, and $|\mathcal{Y}_1| + ... + |\mathcal{Y}_L| = N$. A partition of $\mathcal{Y}$ into the form (1) is widely known as the union-of-subspaces model (Vidal, 2011), which underpins a panoply of practical data clusters ranging from human face images, hand-written digits, to trajectories of moving objects in videos. Given $\mathcal{Y}$ with unknown L and d_k, $1 \leq k \leq L$, the task of subspace clustering is to uncover the partition (1). Among existing solutions to this problem, sparse subspace clustering (SSC) (Liu et al., 2013; Li et al., 2017; Lu et al., 2019; Elhamifar & Vidal, 2013), catalyzed by the witnessed success of compressive sensing (CS) and sparse representation (Baraniuk, 2007; Candès & Wakin, 2008; Davenport et al., 2011; Elad, 2010), has gained much attention because of its compelling experimental performance and provable performance guarantees. Typically, SSC first employs sparse regression for neighbor identification and then constructs a similarity graph of the given dataset, followed by spectral clustering (von Luxburg, 2007) for final data segmentation. In the literature, standard sparse regression for SSC relies on the ℓ_1-minimization approach and its variants thereof (Elhamifar & Vidal, 2013), which are however computationally demanding, especially for large-scale high-dimensional datasets. Accordingly, low-complexity alternatives using greedy-based neighbor selection, e.g., orthogonal matching

pursuit (OMP) (Davenport et al., 2011; Elad, 2010), has been proposed that can perform on par with ℓ_1-minimization in many cases.

In addition to algorithm development for SSC, investigation of their mathematical performance guarantees using fruitful analytical tools from CS also received considerable attention. The vast majority of related works focused on investigating sufficient conditions ensuring the so-called subspace detection property (SDP) (Soltanolkotabi & Candès, 2012), that is, sparse regression certainly returns a neighbor group from the correct cluster, thereby producing a similarity graph without inter-cluster edges (see (Soltanolkotabi & Candès, 2012; Soltanolkotabi et al., 2014; Wu et al., 2021; Wang & Xu, 2016; Wang et al., 2019) regarding the ℓ_1-minimization solutions, and (Dyer et al., 2013; Heckel & Bölcskei, 2015; You et al., 2016b; Tschannen & Bölcskei, 2018) pertinent to greedy search). From the perspective of network connectivity, errorless neighbor identification alone as claimed by SDP is an overly pessimistic condition. This is because, on the one hand, SDP is not necessary for perfect data segmentation. Indeed, as reported in many studies (Ng et al., 2001; Vershynin, 2018), a known type of similarity graphs effectuating successful clustering is one configured with many intra-cluster and few inter-cluster edges. Evidently, this arises when the sparse regression mis-identifies few neighbors, hence violating SDP. The downside of few falsely directed edges from cluster to cluster can be effectively compensated by spectral clustering (see (von Luxburg, 2007) for more discussions on this issue). On the other hand, SDP neither guarantees errorless clustering, especially when accompanied by meager neighbors. The reason is that, despite no inter-cluster edges, the resultant similarity graph is cut into excessively many isolated pieces; poor graph connectivity in this way tends to cause over-estimation of the number L of clusters, leading to a large data clustering error[1] (Soltanolkotabi et al., 2014; Wu et al., 2021). The above facts altogether shed further light on the study of sparse regression for SSC. On aspect of algorithm design, the efforts shall be particularly geared towards fast acquisition of plentiful neighbors, hopefully many true, so that the similarity graph can be fleshed out in a right configuration. As to neighbor recovery performance guarantees, much remains to be explored in quest to new leitmotivs not so stringent as SDP, especially able to reflect neighbor recovery error. Under the framework of two-step weighted ℓ_1-minimization, (Wu et al., 2021) analyzed the neighbor recovery rate, specifically, the probability that the sparse regressor produces at least $k_t(>0)$ correct and at most $k_f(\geq 0)$ incorrect neighbors. A probabilistic characterization of this sort is intuitive and quite flexible, directly taking account of the general case with neighbor mis-identification; when specializing to $k_f = 0$, i.e., errorless neighbor identification, it then reveals how much chance SDP stands, with no less than k_t recovered neighbors.

1.2 Paper Contributions

This paper aims at tackling the aforesaid challenges by revamping the OMP, considering its up-to-par performance, reduced computational complexity and, most importantly, the inherent flexibility to boost neighbor acquisition. Pivoted on the generalized orthogonal matching pursuit (GOMP), a prominent variant of OMP that is first introduced in the literature of CS (Wang et al., 2012) and allowed to identify multiple neighbors per iteration, we propose a new sparse regression scheme for SSC, together with an in-depth analytic study of its neighbor recovery performance guarantee. Specific technical contributions of this paper are summarized as follows.

- We propose to employ GOMP as an effective alternative to OMP for fast neighbor identification. In particular, we first point out the deviation (caused by noise) of the residual vector from the desired ground truth subspace, pinned down by the Angle of Deviation (AoD), plays a pivotal role in neighbor identification. According to this fact, we then argue that GOMP, while digging out more neighbors per iteration, enjoys a smaller AoD, which makes it more resilient to noise corruption and able to achieve higher neighbor identification accuracy than OMP.
- Efficient stopping rules are crucial for greedy-based neighbor selection. If the algorithm stops early, we would end up with scant correct neighbors, yet if late, with overly many false ones; either case is apt to cause poor graph connectivity and eventual erroneous data clustering. Alongside the proposed GOMP we devise a new stopping rule geared toward fulfilling the dimension-aware

[1] Once a similarity graph is constructed, e.g., using either ℓ_1-minimization or greedy methods, one can further employ pruning schemes such as Yang et al. (2020); Qin et al. (2023) to obtain an updated graph with improved connectivity.

property, that is, the algorithm is halted once the number of recovered neighbors is fairly close to the subspace dimension (this is in general a right moment to leave off neighbor search since the residual thereafter is typically dominated by noise). Mathematically, the proposed stopping rule judges the ratio of residual norms over consecutive two iterations against a threshold dependent on the ambient space dimension n, which is known once the dataset is given, and the number p of neighbors identified per iteration that is at the designer's disposal. Advantageously, this dispenses with extra off-line estimation of the subspace dimension or noise strength as required in the existing solutions (Soltanolkotabi & Candès, 2012; Soltanolkotabi et al., 2014; Wu et al., 2021; Wang & Xu, 2016; Wang et al., 2019; Dyer et al., 2013; Heckel & Bölcskei, 2015; You et al., 2016b; Tschannen & Bölcskei, 2018).

- Capitalized on the work (Wu et al., 2021) and under the semi-random model (Soltanolkotabi & Candès, 2012; Soltanolkotabi et al., 2014), we conduct recovery rate analysis to bear out the claimed merits of the proposed GOMP scheme. Supposing that the ground truth subspaces are well-separated from each other, we first derive an analytic probability lower bound for the event that at least k_m neighbors ($0 \leq k_m \leq p$) are correct (i.e., from the ground truth subspace) in the mth iteration. Such a local iteration-wise recovery rate result is then exploited to obtain the global recovery rate, namely, the probability lower bound for the event that at least k_t correct neighbors in total are identified throughout. The obtained analytic formula shows that, for a large data size N and small noise power, GOMP enjoys a higher correct neighbor recovery rate than the conventional OMP, confirming GOMP can facilitate fast acquisition of many correct neighbors. Finally, we prove that, with a high probability, the proposed stopping rule possesses the dimension-aware property.

- To implement the proposed GOMP method, the number p of neighbors identified per iteration should be set beforehand. By further analyzing the obtained recovery rate formulae, the impact of the parameter p on the recovery rate is first discussed. For real-world datasets, oftentimes unable to meet the assumptions required by the semi-random model, recovery rate analysis is rather difficult to carry through. We therefore conduct numerical simulations to investigate the selection of p aimed at fulfilling the dimension-aware property. Interestingly, both the recovery rate analysis for the semi-random model and our simulation study for real-world datasets indicate that a large p is preferred when (i) the ground truth subspaces are well-separated from each other, or (ii) the data size N is large. Some guidelines for selecting p are suggested accordingly.

1.3 Connection to Previous Works

Efficient sparse regression for neighbor identification has played a pivotal role for SSC. The standard ℓ_1-minimization based method was first introduced in the landmark paper (Elhamifar & Vidal, 2013), and since then many related solutions have been proposed. In (Soltanolkotabi et al., 2014; Wu et al., 2021), iterative re-weighted ℓ_1-minimization was adopted to further improve neighbor identification accuracy at the expense of higher algorithmic complexity. In (You et al., 2018; Peng et al., 2013; Matsushima & Brbic, 2019), computationally-efficient solutions under the framework of ℓ_1-minimization were then proposed; the basic idea therein was to pre-process a small amount of data points to acquire side information about the ground truth subspaces, based on which a pruned dataset can then be used to reduce computations. Notably, (You et al., 2016a; Panagakis & Kotropoulos, 2014) utilized mixed-norm regularization to further enhance connectivity of the similarity graph. In contrast to ℓ_1-minimization, greedy search such as OMP is one widely considered low-complexity neighbor identification scheme (Dyer et al., 2013; Heckel & Bölcskei, 2015; You et al., 2016b; Tschannen & Bölcskei, 2018). Recently in (Chen et al., 2018; Zhu et al., 2019), modified OMP algorithms aimed at improving network connectivity have also been proposed; the methods therein utilized the already established neighbor connections to narrow down the candidate neighbor list, overall promoting neighbor recovery and consequently better network connectivity. It is worthy of noting that all the existing OMP-based solutions identify one neighbor per iteration, and employ stopping rules calling for a knowledge of the subspace dimension or noise strength. Boosting neighbor recovery via multiple neighbor identification per iteration, as well as development of efficient stopping rules free from aforementioned side information, is not yet addressed in the literature of SSC.

Regarding the study of mathematical performance guarantees, sufficient conditions for SDP under the ℓ_1-minimization framework have been investigated in, say, (Elhamifar & Vidal, 2013; Soltanolkotabi & Candès, 2012) for the noiseless case, and (Soltanolkotabi et al., 2014; Wang & Xu, 2016; Wang et al., 2019) for the noisy case. Elhamifar & Vidal (2013) utilized convex geometry techniques to derive sufficient conditions tailored for specialized subspace orientations (e.g., disjoint or independent subspaces), while Soltanolkotabi & Candès (2012) dealt with the generalization to subspaces with non-trivial intersection. Under noise corruption, Soltanolkotabi et al. (2014) leveraged certain approximation of the LASSO functional and the restricted isometry property of the noisy data matrix to estimate the probability that SDP holds; for LASSO sparse regression, Wang & Xu (2016) further investigated sufficient conditions that the regularization parameter must satisfy in order to guarantee SDP. Recently, Wu et al. (2021) extended the study in (Soltanolkotabi et al., 2014) to provide recovery rate analyses for general neighbor recovery events. For SSC employing greedy neighbor identification, Dyer et al. (2013) and You et al. (2016b) considered the noiseless scenario and derived sufficient conditions for SDP using convex geometry analysis; Tschannen & Bölcskei (2018) then extended the results in (Dyer et al., 2013) and (You et al., 2016b) to the Gaussian-noise setting, and derived probability lower bounds for the event the SDP holds. As far as we can see, all existing studies of performance guarantees for OMP-based SSC revolved around the fulfillment of SDP; the general case when neighbor mis-identification occurs is left unaddressed.

To sum up, while GOMP has been investigated in CS (Wang et al., 2012), its application and potential impacts to SSC remain yet to be explored. This paper is a first step toward this goal. Thanks to multiple neighbor identification per iteration, the proposed GOMP method boosts neighbor recovery at lower algorithmic complexity as compared to conventional OMP. In addition, our newly developed stopping rule enjoys the dimension-aware property, thereby free from off-line subspace dimension estimation. Moreover, we leverage recovery rate analysis to derive mathematical performance guarantees for general neighbor recovery events. In view of the above points, our study of SSC with GOMP can contribute to more well-rounded literature of SSC under the framework of greedy neighbor selection.

The rest of this paper is organized as follows. Section 2 first explains why GOMP can outdo conventional OMP, and introduces the foundations behind the proposed stopping rule; after that, algorithmic complexity of OMP and the proposed GOMP are analyzed. Section 3 presents the recovery rate analyses, and then discusses the issue of selecting the number p of recovered neighbors per iteration. Section 4 provides numerical simulations to verify our theoretical study and discussions in Section 3. Section 5 goes through the proofs of the main mathematical results. Finally, Section 6 concludes this paper. To ease reading, some detailed technical proofs are relegated to the appendix.

2 Proposed SSC-GOMP

This section introduces the proposed scheme. We first brief in Section 2.1 the reason why multiple neighbor recovery in each iteration is favored, in an attempt to motivate our GOMP proposal. In Section 2.2, we then encapsulate the foundations behind the proposed stopping rule. Finally, in Section 2.3 we provide algorithmic complexity analysis to justify the computational efficiency of the proposed GOMP as compared to OMP.

2.1 Why Multiple Neighbor Recovery per Iteration?

Recall that OMP iteratively identifies a neighbor each time as the data point when paired with the residual vector yields peak absolute inner product. Hence, the orientation of the residual vector in each iteration, in particular, the degree to which it deviates from the ground truth subspace, is important for accurate neighbor identification. To formalize this notion, assume that we are building a neighbor list for the data point $\mathbf{y}_i$ coming from the cluster $\mathcal{Y}_k$ whose ground truth subspace is $\mathcal{S}_k$. Impaired by noise, the residual vector $\mathbf{r}_m^{(i)}$ computed in the mth iteration ($m \geq 1$) is perturbed outwards $\mathcal{S}_k$. If we write $\mathbf{r}_m^{(i)} = \mathbf{r}_{m,\|}^{(i)} + \mathbf{r}_{m,\perp}^{(i)}$, where $\mathbf{r}_{m,\|}^{(i)} \in \mathcal{S}_k$ and $\mathbf{r}_{m,\perp}^{(i)} \in \mathcal{S}_k^{\perp}$ (the orthogonal complement of $\mathcal{S}_k$), such perturbation can be pinned down by the angle of deviation (AoD)

$$\phi_m^i \triangleq \tan^{-1}(\|\mathbf{r}_{m,\perp}^{(i)}\|_2 / \|\mathbf{r}_{m,\|}^{(i)}\|_2) \tag{2}$$

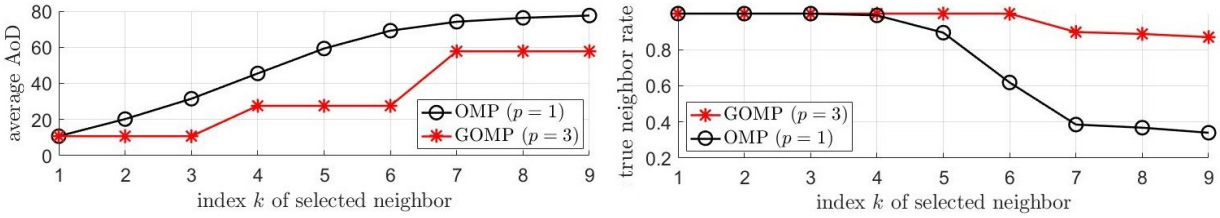

Figure 1: Comparison of GOMP and OMP in terms of average AoD and empirical recovery rate. We consider a synthetic data set of 135 vectors drawn from $L = 3$ orthogonal subspaces, each of a dimension 9, in an ambient domain $\mathbb{R}^{100}$; 45 data points per cluster. The data vectors are sampled uniformly from the intersection of the unit-sphere in $\mathbb{R}^{100}$ with the ground truth subspace, and are corrupted by zero mean Gaussian noise with variance 0.04. For GOMP, $p = 3$ neighbors are picked per iteration as those when matched to the residual yielding the largest three absolute inner products. A total number of 9 neighbors are recovered using both OMP and GOMP. (a) Left: plot of average AoD upon detection of the kth neighbor, $1 \leq k \leq 9$. For GOMP, every three neighbors are detected in each iteration based on the same residual, leading to a staircase AoD curve. Clearly, GOMP results in smaller AoD thanks to fewer iterations. (b) Right: plot of the true neighbor rate, i.e., the fraction of true neighbors recovered, versus the index k of the detected neighbor. Benefiting from smaller AoD, GOMP is seen to improve neighbor identification accuracy.

whereby a large ϕ_m^i means $\mathbf{r}_m^{(i)}$ severely deviates from $\mathcal{S}_k$. As the OMP algorithm iterates, perturbation of the residual would become increasingly severe. This is mainly because, owing to orthogonal projection (see step 3 of Algorithm 1 at the next page), the current residual $\mathbf{r}_m^{(i)}$ is obtained by removing from from the previous $\mathbf{r}_{m-1}^{(i)}$ the component lying in the subspace spanned by the already-selected neighbors, most of which are likely correct. Consequently, the magnitude $\|\mathbf{r}_{m,\|}^{(i)}\|_2$ of the component $\mathbf{r}_{m,\|}^{(i)} \in \mathcal{S}_k$ diminishes from iteration to iteration; instead, the term $\|\mathbf{r}_{m,\perp}^{(i)}\|_2$, which reflects the strength of projected mis-identified neighbors (if any) plus noise onto $\mathcal{S}_k^{\perp}$, is typically non-decreasing with m. Put together, the cascade effect is therefore an increase in ϕ_m^i with m, rendering the residual $\mathbf{r}_m^{(i)}$ more and more prone to neighbor mis-identification as the algorithm iterates. In this regard, a simple remedy for securing enough correct neighbors in few iterations (a "small m" is favored) is therefore to identify multiple neighbors per iteration, say, the p data points ($p > 1$) corresponding to the largest p absolute inner products. GOMP is therefore a potential solution to meet this goal. Using synthetic data, Fig. 1-(a) clearly demonstrates GOMP yields smaller average AoD than OMP thanks to fewer iterations[2]; this accordingly brings about higher neighbor identification accuracy, as illustrated in Fig. 1-(b) (this issue will be elaborated in Section 3).

The reduction in the number of iterations can moreover reduce algorithmic complexity, which is potentially appealing in real-time applications. Indeed, the major computational bottleneck of the OMP algorithm is the orthogonal projection operation (step 3 in Algorithm 1, with $p = 1$). To recover $p(> 1)$ neighbors, conventional OMP requires p iterations, hence p orthogonal projections. Instead, GOMP calls for just one iteration, so one orthogonal projection only, and is therefore more computationally efficient (detailed algorithmic complexity comparison of OMP and GOMP is given at the end of this section). Considering all the above facts, we thus propose to adopt GOMP in place of OMP for neighbor identification.

2.2 Halt When There Are About as Many Neighbors as Subspace Dimension

Since the data point $\mathbf{y}_i$ is associated with the ground truth subspace $\mathcal{S}_k$, a group of around d_k true neighbors would reach a "critical mass" to well explain $\mathbf{y}_i$ and, if so, the residual from then on is highly apt to be dominated by noise, standing very little chance to uncover more true neighbors (this will be born out by our mathematical analysis in Section 3). Grounded on this fact, the algorithm is expected to be halted once d_k neighbors or so are available. We shall recall the stopping rule widely adopted in the literature (You et al., 2016b; Tschannen & Bölcskei, 2018), which terminates neighbor search when either the number of iterations

[2]An analytic study of average AoD for GOMP/OMP is rather challenging, and is one of our future works.

reaches a pre-set bound M, or the residual becomes so small that $\|\mathbf{r}_m^{(i)}\|_2 \leq \tau$ for some threshold $\tau > 0$. Though implicit, this assumes availability of prior knowledge about M and τ; the former is arguably all about the subspace dimension d_k and the latter is closely related to the background noise strength, either of which can only be acquired through extra off-line estimation process. Considering that the dedicated overhead of parameter estimation could be costly, below we develop a new stopping rule which is per se aware of the subspace dimension without the need of knowing M or τ.

To introduce the proposed approach, let us write the data point under consideration as $\mathbf{y}_i = \mathbf{x}_i + \mathbf{e}_i$, where $\mathbf{x}_i$ is the noiseless signal point and $\mathbf{e}_i$ is the additive noise. The residual vector $\mathbf{r}_m^{(i)}$, which can be obtained from $\mathbf{y}_i$ through a sequence of m orthogonal projections, can be expressed as

$$\mathbf{r}_m^{(i)} = \prod_{l=1}^{m} \mathbf{P}_l \mathbf{y}_i = \prod_{l=1}^{m} \mathbf{P}_l (\mathbf{x}_i + \mathbf{e}_i) = \prod_{l=1}^{m} \mathbf{P}_l \mathbf{x}_i + \prod_{l=1}^{m} \mathbf{P}_l \mathbf{e}_i, \tag{3}$$

where $\mathbf{P}_l$ is the orthogonal projection onto the orthogonal complement of the subspace spanned by the already-selected neighbors up to the lth iteration. Assume that most of the recovered neighbors up to the $m-1$ iterations are from the correct subspace $\mathcal{S}_k$; as such, the projected signal is very small and the residual $\mathbf{r}_m^{(i)}$ is strongly dominated by the projected noise so that $\mathbf{r}_m^{(i)} \approx \prod_{l=1}^{m} \mathbf{P}_l \mathbf{e}_i$. This is typical the case once about as many neighbors as the subspace dimension are recovered. If the noise $\mathbf{e}_i$ is Gaussian, so is the residual $\mathbf{r}_m^{(i)}$, which, being nearly isotropic, tends to distribute its power evenly over all the dimensions (about $n-d_k$) of the orthogonal complement of the subspace spanned by all the already-selected neighbors; that is to say, each dimension shares a factor $1/(n-d_k)$ of the total power $\|\mathbf{r}_m^{(i)}\|_2^2$. During the $m+1$th iteration, $\mathbf{r}_{m+1}^{(i)}$ is then obtained from $\mathbf{r}_m^{(i)}$ by removing from it the components along the p newly selected neighbors, implying that $\|\mathbf{r}_m^{(i)} - \mathbf{r}_{m+1}^{(i)}\|_2^2$ is close to $p \times \|\mathbf{r}_m^{(i)}\|_2^2/(n-d_k) \approx p \times \|\mathbf{r}_m^{(i)}\|_2^2/n$, in which the approximation makes sense since the subspace dimension d_k is in general very small in comparison with the ambient space dimension n. Taking square root and using the triangle inequality, we then obtain the following condition to halt neighbor search

$$\|\mathbf{r}_m^{(i)}\|_2 - \|\mathbf{r}_{m+1}^{(i)}\|_2 \leq (\|\mathbf{r}_m^{(i)}\|_2 \sqrt{p})/\sqrt{n}, \tag{4}$$

or equivalently,

$$\frac{\|\mathbf{r}_{m+1}^{(i)}\|_2}{\|\mathbf{r}_m^{(i)}\|_2} \geq 1 - \sqrt{p/n}. \tag{5}$$

The proposed halting rule (5) is dimension-aware because, for most cases, it is triggered once the residual $\mathbf{r}_m^{(i)}$ is dominated by noise owing to the recovery of d_k neighbors or thereabouts. We should moreover note that the left-hand-side of (5) admits the form of a residual norm ratio, which advantageously rids off the knowledge of noise strength. Indeed, since $\mathbf{r}_m^{(i)} \approx \prod_{l=1}^{m} \mathbf{P}_l \mathbf{e}_i$, we have

$$\|\mathbf{r}_m^{(i)}\|_2 \approx \|\prod_{l=1}^{m} \mathbf{P}_l \mathbf{e}_i\|_2 = \alpha_m \|\mathbf{e}_i\|_2, \text{ for some } 0 < \alpha_m < 1. \tag{6}$$

The condition (6) motivates the stopping rule of the form You et al. (2016b); Tschannen & Bölcskei (2018), whereby the threshold τ depends on the noise strength $\|\mathbf{e}_i\|_2$. Using (6), the residual norm ratio in the proposed stopping rule (5) reads

$$\frac{\|\mathbf{r}_{m+1}^{(i)}\|_2}{\|\mathbf{r}_m^{(i)}\|_2} \approx \frac{\alpha_{m+1}}{\alpha_m}, \text{ for some } 0 < \alpha_{m+1} < \alpha_m < 1. \tag{7}$$

which is clearly independent of the noise strength $\|\mathbf{e}_i\|_2$. The decision threshold in (5) assumes nothing more than a knowledge of the ambient dimension n, which is always known in advance, and the parameter p, which is at the designer's disposal. It is therefore free from the need of extra off-line estimation of the subspace dimension or noise strength, in marked contrast with the existing solutions You et al. (2016b); Tschannen & Bölcskei (2018). We summarize the proposed SSC-GOMP algorithm in Algorithm 1.

Algorithm 1 SSC-GOMP algorithm with the proposed data dependent stopping criterion

Input: Observed data set $\mathcal{Y} = \{\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N\}$
for $i = 1$ **to** N **do**
Let $m = 0$, $\mathbf{r}_0 = \mathbf{y}_i$, $\mathbf{r}_{-1} = 2\mathbf{y}_i$, $\Lambda_0 = \phi$.
if $(1 - \|\mathbf{r}_m^{(i)}\|_2 / \|\mathbf{r}_{m-1}^{(i)}\|_2 \geq \sqrt{p/n})$ **then**
1) $m = m + 1$.
2) $\Lambda_m = \Lambda_{m-1} \cup \mathcal{T}_m$, where $\mathcal{T}_m$ is the set of cardinality p such that
$$|\langle \mathbf{y}_j, \mathbf{r}_{m-1} \rangle| \geq |\langle \mathbf{y}_q, \mathbf{r}_{m-1} \rangle|, \forall j \in \mathcal{T}_m,\ q \notin \mathcal{T}_m.$$
3) $\mathbf{r}_m = (\mathbf{I} - \mathbf{Y}_{\Lambda_m} \mathbf{Y}_{\Lambda_m}^{\dagger}) \mathbf{r}_{m-1}$.
end if
4) $\mathbf{c}_i^* = \arg\min_{\mathbf{c}: supp(\mathbf{c}) \subset \Lambda_{m-1}} \|\mathbf{y} - \mathbf{Y}\mathbf{c}\|_2$
5) Normalize the column vector $\mathbf{c}_i^*$ to be unit-norm. Let $\overline{\mathbf{c}}_i^* = [c_{i,1}^* ... c_{i,i-1}^*\ 0\ c_{i,i+1}^* ... c_{i,N}^*]^T \in \mathbb{R}^n$.
end for
6) Set $\mathbf{C} = [c_{i,j}] = [\overline{\mathbf{c}}_1^* \cdots \overline{\mathbf{c}}_N^*]$, and $\mathbf{G} = [g_{i,j}]$, where $g_{i,j} = |c_{i,j}| + |c_{j,i}|$.
7) Form a similarity graph with N nodes, with the weight on the edge between the (i, j) node pair equal to $g_{i,j}$.
8) Apply spectral clustering to the similarity graph.
Output: sparse representation vectors $\{\mathbf{c}_i^*, ..., \mathbf{c}_N^*\}$

2.3 Algorithmic Complexity

We end this section by analyzing algorithmic complexity of the proposed GOMP and OMP. To ease discussion, assume that pM neighbors are to be recovered. Then OMP requires pM iterations, in which the mth iteration has complexity $\mathcal{O}(nN + nm + nm^2 + m^3)$, $1 \leq m \leq pM$, respectively (Sturm & Christensen, 2012). Instead, GOMP requires just M iterations, the mth iteration with complexity $\mathcal{O}(nN + npm + n(pm)^2 + (pm)^3)$, $1 \leq m \leq M$, respectively. Notice that, for $1 \leq m \leq M$, the pmth iteration of OMP has the same complexity as the mth iteration of GOMP. Hence, it is clear that OMP scheme calls for additional $(p-1)M$ iterations, among which the complexity is $\mathcal{O}(nN + nm + nm^2 + m^3)$ for $1 \leq m \leq pM$, $m \neq p, 2p, ..., Mp$. Evidently, GOMP is more computationally efficient than OMP.

3 Theorectical Results

In this section, we present the recovery rate analysis for the proposed SSC-GOMP. Under the semi-random model assumption, Section 3.1 first derives analytic recovery rate formulae. Section 3.2 then discusses the guidelines for selecting the number p of recovered neighbors per iteration.

3.1 Recovery Rate Analysis

To formalize matters, the data vector is assumed to follow the standard additive noise model, that is,

$$\mathbf{y}_i = \mathbf{x}_i + \mathbf{e}_i, 1 \leq i \leq N, \tag{8}$$

where $\mathbf{x}_i \in \mathbb{R}^n$ is the unit-norm noiseless signal vector and $\mathbf{e}_i \in \mathbb{R}^n$ is the noise. The analyses below are built on the popular semi-random model (Soltanolkotabi & Candès, 2012; Soltanolkotabi et al., 2014; Wu et al., 2021; Wang & Xu, 2016), i.e., the ground truth subspaces $\mathcal{S}_1, ..., \mathcal{S}_L$ in the partition (1) are fixed but otherwise unknown, whereas the data vectors and noise are random. Such a model is widely used in the theoretical study of SSC, thanks to its interpretability and amiability to analysis. Similar to the previous works (Wu et al., 2021; Wang et al., 2019; Heckel & Bölcskei, 2015; Tschannen & Bölcskei, 2018), the following assumptions are made in the sequel.

Assumption 1. *For each $1 \leq i \leq N$, the signal vector $\mathbf{x}_i \in \mathbb{R}^n$ is uniformly sampled from $\mathcal{B}_k$, where $\mathcal{B}_k \triangleq \{\mathbf{x} \mid \mathbf{x} \in \mathbb{R}^n, \|\mathbf{x}\|_2 = 1\} \cap \mathcal{S}_k$ is the intersection of the unit sphere with the subspace $\mathcal{S}_k$.* □

Assumption 2. *For each $1 \leq i \leq N$, the noise $\mathbf{e}_i \in \mathbb{R}^n$, $1 \leq i \leq N$, are i.i.d. Gaussian random vectors with zero mean and covariance matrix $(\sigma^2/n)\mathbf{I}$, and are independent of the signal vectors $\mathbf{x}_i$'s.* □

To gauge the degree to which two subspaces are separated away from each other, we recall the affinity between two distinct subspaces $\mathcal{S}_k$ and $\mathcal{S}_l$ that is defined to be (Soltanolkotabi & Candès, 2012)

$$aff(\mathcal{S}_k, \mathcal{S}_l) \triangleq \frac{\|\mathbf{U}_k^T \mathbf{U}_l\|_F}{\sqrt{\min\{d_k, d_l\}}}, \tag{9}$$

where columns of $\mathbf{U}_k$ ($\mathbf{U}_l$, respectively) form an orthonormal basis for $\mathcal{S}_k$ ($\mathcal{S}_l$, respectively).

Assumption 3. *The subspace affinity satisfies*

$$\max_{k,k\neq l} aff(\mathcal{S}_k, \mathcal{S}_l) + \frac{9\sqrt{3} d_L (1+\sigma)}{(8-12\sigma)\sqrt{n-d_L}\log N} \leq \frac{\tau}{4\log N}, \tag{10}$$

where $0 < \tau < 1$. □

Notably, Assumptions 3 guarantees that different subspaces are well separated from each other; affinity conditions akin to (10) are also needed in many existing studies of performance guarantees for SSC (Wu et al., 2021; Wang et al., 2019; Heckel & Bölcskei, 2015; Tschannen & Bölcskei, 2018). Without loss of generality, we assume in the sequel $\mathbf{x}_N \in \mathcal{S}_L$, therefore $\mathbf{y}_N \in \mathcal{Y}_L$, and our goal is to identify a neighbor group of $\mathbf{y}_N$. Under the above three assumptions it can be shown that GOMP stands a high chance to recover many true neighbors in each of the first $\lceil d_L/p \rceil$ iterations. More precisely, we have the following theorem.

Theorem 1. *(Iteration-Wise Recovery Rate): Let $\{k_1, k_2, ..., k_M\}$ be a sequence of integers satisfying $0 \leq k_m \leq p$, for $1 \leq m \leq M \leq \lceil d_L/p \rceil$. Under Assumptions 1 to 3, the proposed SSC-GOMP obtains at least k_m true neighbors at the mth iteration, $1 \leq m \leq M$, with a probability exceeding*

$$\begin{aligned} 1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - p(M-1)} - \sum_{m=1}^{M} \Bigg[\Bigg(\Bigg(\frac{2e(N-|\mathcal{Y}_L|)}{(p-k_m+1)N^{8\log N/d_L}} \Bigg)^{p-k_m+1} \\ + \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-k_m} \left(\frac{e(|\mathcal{Y}_L|-1)}{k_m-1}\right)^{k_m-1} + \frac{4+2c}{N^2} \Bigg) \mathbf{1}(k_m > 0) \Bigg], \end{aligned} \tag{11}$$

where $\mathbf{1}(\bullet)$ is the indicator function, $c > 0$ is a constant, and d_L is the dimension of the subspace $\mathcal{S}_L$.

Proof : See Section 5.1. □

A further scrutiny reveals the lower bound (11) is high in most practical cases. Indeed, with a large ambient dimension n and small noise level σ, the second and third terms in (11) are kept small; regarding the last summation, the first and third terms scale like $N^{[1-(8\log N/d_L)](p-k_m+1)}$ and N^{-2}, respectively, whereas the second term decays exponentially fast as $|\mathcal{Y}_L|$ increases, thanks to $0 < \tau < 1$ (see Assumption 3). When specialized to $p = 1$ and $k_m = 1$ for all $1 \leq m \leq M$, i.e., the case with conventional OMP subject to all detected neighbors being true, the lower bound (11) then reads

$$1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - p(M-1)} - M\left[\frac{2e(N-|\mathcal{Y}_L|)}{N^{8\log N/d_L}} + \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-1} + \frac{4+2c}{N^2}\right]. \tag{12}$$

Notably, a probability lower bound akin to (12) was also reported in (Tschannen & Bölcskei, 2018), which addressed sufficient conditions ensuring correct neighbor recovery. Based on the iteration-wise recovery rate result given in Theorem 1, the following corollary further establishes the global recovery rate, namely, the probability of the event that GOMP succeeds in recovering a specified total number of true neighbors throughout the $M \leq \lceil d_L/p \rceil$ iterations.

Corollary 1. *(Global Recovery Rate): Let $0 \leq k_t \leq pM$ be an integer and write $k_t = Mq_t + r_t$, where $0 \leq r_t \leq M-1$. Under the same setup as in Theorem 1, GOMP can recover at least k_t true neighbors in*

total throughout $M(\leq \lceil d_L/p \rceil)$ *iterations with a probability higher than*

$$\begin{aligned}
&1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - p(M-1)} \\
&- r_t \left[\left(\frac{2e(N-|\mathcal{Y}_L)|}{(p-q_t)N^{8\log N/d_L}}\right)^{p-q_t} + \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-q_t-1}\left(\frac{e(|\mathcal{Y}_L|-1)}{q_t}\right)^{q_t} + \frac{4+2c}{N^2}\right] \\
&-(M-r_t)\left[\left(\frac{2e(N-|\mathcal{Y}_L|)}{(p-q_t+1)N^{8\log N/d_L}}\right)^{p-q_t+1} + \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-q_t}\left(\frac{e(|\mathcal{Y}_L|-1)}{q_t-1}\right)^{q_t-1} + \frac{4+2c}{N^2}\right]\mathbf{1}(q_t>0).
\end{aligned} \tag{13}$$

Proof : See Section 5.2. □

By following our examination of (11), it is easy to check the probability lower bound (13) is high in most practical cases. Moreover, with the aid of (13), GOMP is seen to yield a higher true neighbor recovery rate as compared with the conventional OMP. To better illustrate this, we assume without loss of generality that a group of pM neighbors is to be found, while demanding at least $k_t = kM$ true neighbors, for some $1 < k \leq p$ (that is to say, k/p of the total neighbors are true). Hence, GOMP requires M iterations, and OMP p times more. Under this setting, the lower bound (13) for GOMP reads

$$\begin{aligned}
1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - pM + p} &- M\left(\frac{2e(N-|\mathcal{Y}_L|)}{(p-k+1)N^{8\log N/d_L}}\right)^{p-k+1} \\
- M\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-k}&\left(\frac{e(|\mathcal{Y}_L|-1)}{k-1}\right)^{k-1} - \frac{(4+2c)M}{N^2},
\end{aligned} \tag{14}$$

whereas for OMP the bound becomes

$$1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - pM + 1} - \frac{2kMe(N-|\mathcal{Y}_L|)}{N^{8\log N/d_L}} - kM\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-1} - \frac{(4+2c)kM}{N^2}. \tag{15}$$

Note that (14) differs from (15) in the last four terms, among which it is clear that $6\left(\sigma/\sqrt{\pi}\right)^{d_L-pM+p} < 6\left(\sigma/\sqrt{\pi}\right)^{d_L-pM+1}$ for small noise level σ, and $\frac{(4+2c)M}{N^2} < \frac{(4+2c)kM}{N^2}$ because $k > 1$. Since the 4th term $M\left(\frac{2e(N-|\mathcal{Y}_L|)}{(p-k+1)N^{8\log N/d_L}}\right)^{p-k+1}$ in (14) and the 4th term $\frac{2kMe(N-|\mathcal{Y}_L|)}{N^{8\log N/d_L}}$ in (15) scale like $N^{[1-(8\log N/d_L)](p-k+1)}$ and $N^{[1-(8\log N/d_L)]}$, respectively, both vanish whenever the data size N is very large. Also, since $\tau < 1$, both the 5th term $M\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-k}\left(\frac{e(|\mathcal{Y}_L|-1)}{k-1}\right)^{k-1}$ in (14) and the 5th term $kM\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-1}$ in (15) decay exponentially fast to zero as the cluster size $|\mathcal{Y}_L|$ grows, therefore negligible. Accordingly, when the subspaces are well separated from each other and the data/cluster size is large enough, GOMP improves the true neighbor recovery rate.

Theorem 1 and Corollary 1 specify the neighbor recovery rate up to the $\lceil d_L/p \rceil$th iteration, upon which GOMP is expected to be halted by the proposed stopping rule (5) thanks to availability of about d_L recovered neighbors. Such a conjecture is provably true with a high probability, as established in the next theorem.

Theorem 2. *Under the same setup as in Theorem 1, GOMP accompanied by the proposed stopping rule (5) halts neighbor search till the* $\lceil d_L/p \rceil$*th iteration with a probability exceeding*

$$\begin{aligned}
&1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - p(\lceil d_L/p \rceil - 1)} - 2pe^{-\sqrt{n/p}} \\
&- \left\lceil \frac{d_L}{p} \right\rceil \left[\frac{2e(N-|\mathcal{Y}_L|)}{N^{8\log N/d_L}} + \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-p}\left(\frac{e(|\mathcal{Y}_L|-1)}{p-1}\right)^{p-1} + \frac{4+2c}{N^2}\right].
\end{aligned} \tag{16}$$

Proof : See Section 5.3. □

Under Assumptions 1 to 3, Theorem 2 provides an analytic probability lower bound for the event that the proposed stopping rule (5) is aware of the subspace dimension, i.e., it terminates the GOMP algorithm when the number of recovered neighbors is close to the subspace dimension d_L. We further remark that, when the noise power increases slightly further, the dimension of the span of data points in $\mathcal{Y}_L$ would exceed d_L; accordingly, the algorithm would stop with more than $\lceil d_L/p \rceil$ iterations in order to recover more neighbors commensurate with the increased dimension. As the noise power grows higher, the residual is soon dominated by noise so as to trigger the stopping condition (5), rendering the algorithm terminated in fewer than $\lceil d_L/p \rceil$ iterations with scant neighbors. This phenomenon will be seen in our simulation study.

3.2 On Selection of p

The performance of the proposed GOMP algorithm depends on the parameter p, i.e., the number of identified neighbors per iteration. With the aid of the recovery rate analyses for the semi-random model, below we first discuss the selection of p aimed at improving the recovery rate. The results will shed light on the selection of p for practical datasets.

Recall the assertion of Theorem 2 that, with a high probability, the proposed stopping rule (5) halts the algorithm once d_L neighbors or so are recovered and, if so, the GOMP algorithm yields a high global recovery rate throughout all iterations according to Corollary 1. Hence, a natural criterion for selecting p is to maximize both the probability lower bounds (13) and (16); in this way, the proposed algorithm stands a high chance of recovering many true neighbors and only few false neighbors. However, we should note that both probability lower bounds are rather complicated functions in p and the ground truth parameters (such as the data size N, subspace dimension d_L, cluster size $|\mathcal{Y}_L|$, noise power σ, and the subspace affinity τ). An explicit and tractable rule for choosing p toward recovery rate enhancement is therefore very difficult to obtain. Even if such a solution can be found, it necessarily depends on the ground truth parameters, which are nonetheless unknown to the designers. Despite this, by analyzing the lower bounds (13) and (16) we can still summarize certain interesting properties of the recovery rate as the parameter p varies.

For this, let us examine each individual term in the probability lower bounds (13) and (16) to see how they change with p. Suppose that a total number $M_t = pM$ of neighbors are recovered, among which at least $k_t (> 0)$ are true. To ease discussions, we assume that $k_t \geq M$ so that $q_t = \lfloor k_t/M \rfloor > 0$, where $\lfloor \bullet \rfloor$ is the floor function; hence, on average at least q_t true neighbors are found in each iteration. We can first observe from (13) the followings:

a) Clearly, the 3rd term $-6(\sigma/\sqrt{\pi})^{d_L - M_t + p}$ in the lower bound (13) increases with p when the noise level is so small that $\sigma < \sqrt{\pi}$.

b) Following the discussions below (15), the 4th and 7th terms of (13) scale like $-N^{[1-(8\log N/d_L)](p-q_t)}$ and $-N^{[1-(8\log N/d_L)](p-q_t+1)}$, respectively. Since

$$p - q_t = p - \lfloor k_t/(M_t/p) \rfloor = p - \lfloor p(k_t/M_t) \rfloor, \tag{17}$$

which increases with p, we conclude that the 4th and 7th terms of (13) also increases with p.

c) To check the 5th and 8th terms of (13), we first note that (i) the subspace affinity bound $\tau < 1$ (see (10)), thereby $\tau\sqrt{2/\pi} < 1$, and (ii) since the number k_t of recovered true neighbors never exceeds the cluster size $|\mathcal{Y}_L|$, we must have $|\mathcal{Y}_L| > k_t > q_t = \lfloor k_t/M \rfloor$, leading to $e(|\mathcal{Y}_L| - 1)/q_t > 1$. Hence, it can be readily seen that the 5th term $-r_t \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L| - d_L - q_t - 1} \left(\frac{e(|\mathcal{Y}_L|-1)}{q_t}\right)^{q_t}$ and the 8th term $-(M - r_t)\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L| - d_L - q_t} \left(\frac{e(|\mathcal{Y}_L|-1)}{q_t - 1}\right)^{q_t - 1}$ decrease with q_t. Notably, since $q_t = \lfloor p(k_t/M_t) \rfloor$ grows as p increases, we therefore conclude that the 5th and 8th terms decrease with p.

d) Finally, the sum of the 6th and 9th terms of (13) equals $\frac{-M(4+2c)}{N^2} = \frac{-M_t(4+2c)}{pN^2}$, which increases with p.

Combining (a)∼(d), we know that the 3rd, 4th, 6th, 7th, and 9th terms of (13) increase with p, whereas the 5th and 8th terms decrease with p. In particular, if the subspace affinity τ is very small, the 5th and 8th terms in (13) can be neglected so that the lower bound (13) increases with p. Hence, as long as the ground truth subspaces are well-separated from each other, a large p is preferred. On the other hand, we consider the scenario that the data size N and cluster size $|\mathcal{Y}_L|$ are large; therefore the 4th, 5th, 7th, and 8th terms in the lower bound (13) are vanishingly small (see also the discussions below (15)). The lower bound (13) is then reduced to

$$1 - Ne^{-n/8} - 6\left(\frac{\sigma}{\sqrt{\pi}}\right)^{d_L - M_t + p} - \frac{M_t(4+2c)}{pN^2}. \tag{18}$$

Clearly, with M_t, N, d_L, and σ being fixed, the lower bound (18) increases with p. Hence, we can also conclude that, as long as the data size N and cluster size $|\mathcal{Y}_L|$ are large enough, a large p is preferred. Next, we consider the lower bound (16). To ease discussion, we assume that p divides d_L. Based on our analyses on the lower bound (13) (see the discussions (a) to (d) above) we note that:

e) The 3rd term $-6(\sigma/\sqrt{\pi})^{d_L - p(\lceil d_L/p \rceil - 1)} = -6(\sigma/\sqrt{\pi})^p$ increases with p when the noise level is so small that $\sigma < \sqrt{\pi}$.

f) The 4th term $-2pe^{-\sqrt{n/p}}$ is negligible since the ambient dimension $n \gg p$.

g) Clearly, the 5th term $-\left\lceil \frac{d_L}{p} \right\rceil \frac{2e(N - |\mathcal{Y}_L|)}{N^{8 \log N/d_L}}$ increases with p.

h) Since $\tau\sqrt{2/\pi} < 1$ and $e(|\mathcal{Y}_L| - 1)/p > 1$ (see discussion (c)), $\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L| - d_L - p} \left(\frac{e(|\mathcal{Y}_L|-1)}{p-1}\right)^{p-1}$ increases with p; obviously $\lceil d_L/p \rceil$ decreases with p. Hence, it is hard to tell whether the 6th term $-\lceil \frac{d_L}{p} \rceil \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L| - d_L - p} \left(\frac{e(|\mathcal{Y}_L|-1)}{p-1}\right)^{p-1}$ of (16) increases with p or not.

i) Finally, the 7th term $-\lceil \frac{d_L}{p} \rceil \frac{4+2c}{N^2}$ increases with p.

By (e), (f), (g), and (i), we know that the 3rd, 4th, 5th, 7th terms of (16) increase with p. Still, if the subspace affinity τ is very small, the 6th term of (16) is negligible, leading to the conclusion that (16) increases with p. Hence, a large p is favored for well-separated ground truth subspaces. On the other hand, when N and $|\mathcal{Y}_L|$ are large, the 5th and 6th terms are vanishingly small (again, see the discussions below (15)) so that (16) increases with p. As a result, a large p is preferred whenever the data size N and cluster size $|\mathcal{Y}_L|$ are large. It is worth noting that the inferred tendency of (16) as p varies totally agrees with that of (13). Finally, we should note that the cluster size $|\mathcal{Y}_L|$ is unknown beforehand. For balanced datasets, wherein all clusters are about equally large, $|\mathcal{Y}_L|$ is typically in direct proportion to the data size N; in this case, we then would like to increase p for large-scale datasets. We can therefore conclude that, for well-separated subspaces (small affinity) or large-scale balanced datasets (large N and $|\mathcal{Y}_L|$), a large p is preferred (the same tendency is also seen in our simulation study using real-world datasets). This is expected since, in the former case, the data point under consideration is surrounded by many true neighbors, and for the latter case tends to be strongly correlated with many neighbors. Either way, the residual would quickly diminish in just few iterations, and one should therefore adopt a large p to boost neighbor recovery before the residual is depleted of the component from the desired subspace.

We should note however that our recovery rate analyses are developed under the semi-random model, that is, the signal points are uniform (Assumption 1) and noise is Gaussian (Assumption 2), which are seldom met in practice. Hence, for real-world datasets, performance guarantees similar to the semi-random model can hardly be obtained. Without such analytic metrics, a simple and intuitively reasonable way of judging a good p would be the fulfillment of the dimension-aware property; this is because availability of about d_L neighbors is likely to begat a similarity graph with good connectivity, i.e., sufficiently many intra-cluster edges and few inter-cluster edges. Based on our simulation study (see Section 4 for more details) it is observed that the dimension-aware property tends to hold for a wide range of p when datasets enjoy a small subspace affinity, such as the Coil-100 (Nene et al., 1996) and MNST datasets (LeCun et al., 1998), but seems

to fail for datasets with medium-to-large subspace affinity, such as the Devanagari (Acharya et al., 2015) and Extended Yale B (Lee et al., 2005) datasets. In particular, the Devanagari dataset is subject to large noise corruption and prefers a large p; contrarily, the Extended Yale B is under small noise corruption and prefers a small p. We would therefore conclude that the dimension-aware property holds when the subspace affinity is small; this is because there would then exist ample intra-cluster neighbors for the data point under consideration, making it highly likely to find about d_L correct neighbors no matter how many are recovered in each iteration. The above simulation findings also suggest that selection of p is highly dependent on the affinity and noise power, which are nonetheless known unless additional off-line estimation is further conducted. In the absence of such side information, we would then propose to select a large p when the data size N is large; this is justified by our simulation results and also agrees with our previous claim in the semi-random model case. In particular, we would suggest choosing $p = 2, 3$ for small-scale datasets such as Extended Yale B, Coil-100, and NCKU human face (Chen & Lien, 2009) datasets (all with thousands of data points or less), while $p = 5, 6$ for large-scale datasets, e.g. MNIST, Devanagari, and Cifar-10 (Krizhevsky, 2009) datasets (with tens of thousands of data points or more).

Remark: In the context of sparse signal recovery via GOMP, selection of p in general depends on the sparsity level, but the development of explicit rules for choosing p still remains a difficult open problem (Fu et al., 2022).

4 Experimental Results

In this section, numerical simulations based on both synthetic data and real-world data are used to corroborate our theoretical study and illustrate the performance of the proposed GOMP method. Synthetic data are gen-erated similar to Elhamifar & Vidal (2013); Wu et al. (2021). The ground truth is a union of three subspaces $\mathcal{S}_1$, $\mathcal{S}_2$ and $\mathcal{S}_3$ of $\mathbb{R}^{350}$, with an equal subspace dimension $d_1 = d_2 = d_3 = 6$; separation be-tween two distinct subspaces is gauged by the subspace affinity defined in (9). Noiseless signal points are uniformly drawn at random from each subspace, and are corrupted by Gaussian noise with zero mean and standard deviation σ. The sampling density is defined to be $\phi_k \triangleq |\mathcal{Y}_k|/d_k$, $1 \leq k \leq 3$. Regarding the neighbor identification performance metrics, we consider the true neighbor rate (TNR) (Wu et al., 2021), that is,

$$TNR \triangleq \frac{\sum_{(i,j)\in\mathcal{T}} \mathbf{1}(c^*_{i,j} \neq 0)}{\sum_{1\leq i,j\leq N} \mathbf{1}(c^*_{i,j} \neq 0)},\ \mathcal{T} \triangleq \{(i,j)|\mathbf{y}_i, \mathbf{y}_j \in \mathcal{Y}_l,\ \text{for some } 1 \leq l \leq L\}, \tag{19}$$

where $c^*_{i,j}$ is the jth entry of the computed sparse representation vector $\mathbf{c}^*_i$ using GOMP/OMP (see step 4 of the algorithm), and the average number of recovered neighbors (ANRN)

$$ANRN \triangleq \sum_{i=1}^{N} \|\mathbf{c}^*_i\|_0/N, \tag{20}$$

which assesses the ability to boost neighbor acquisition. As in (Soltanolkotabi & Candès, 2012; Soltanolkotabi et al., 2014; Wu et al., 2021; You et al., 2018; Matsushima & Brbic, 2019), the global clustering performance is evaluated based on the correct clustering rate (CCR), defined as beginequation

$$CCR \triangleq (\#\text{ of correctly clustered data points})/N. \tag{21}$$

4.1 Synthetic Data

We first use synthetic data to test the performance of the proposed method. To ease illustration, we consider the case that $aff(\mathcal{S}_k, \mathcal{S}_l) = \rho$ for all $1 \leq k \neq l \leq 3$, i.e., the three subspaces are equally separated from each other, and that $\phi_1 = \phi_2 = \phi_3 = \phi$, so with equal sample density. Associated with 9 different pairs (ρ, ϕ) of subspace affinity and sample density, Fig. 2 and 3 plot the simulated TNR and CCR, respectively, versus noise standard deviation σ; sub-figures on the same row (column, respectively) correspond to an identical ϕ (ρ, respectively), while ρ (ϕ, respectively) is increased when going from left to right (top to bottom, respectively). In generating Fig. 2 and 3, the total number of iterations is pre-set to be $M = 6$ for OMP

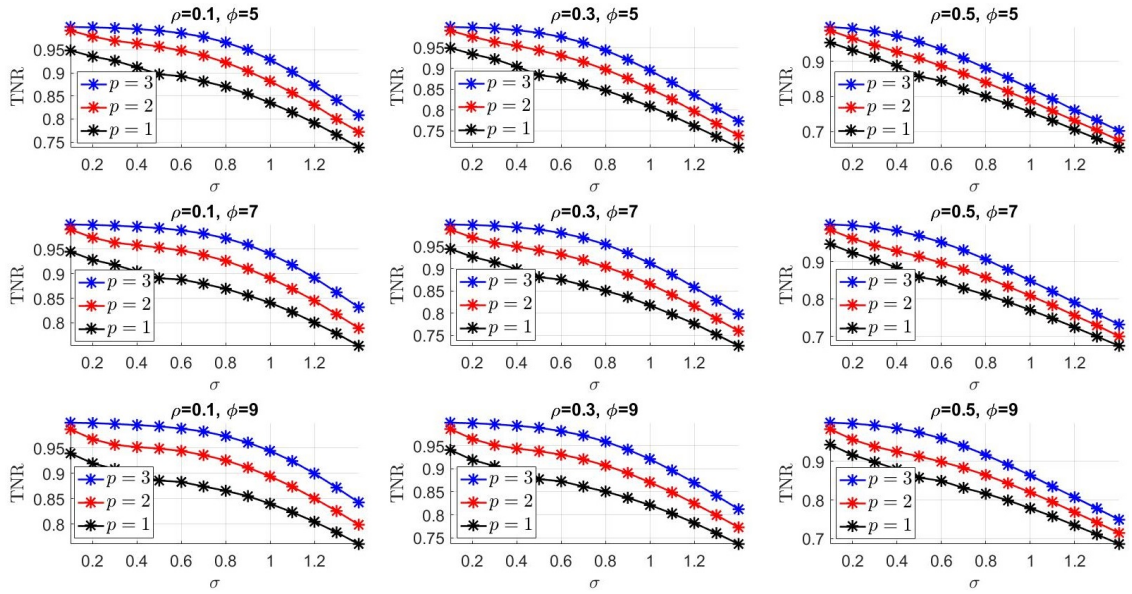

Figure 2: TNR versus noise standard deviation σ for nine different pairs of subspace affinity ρ and sample density ϕ; assume the subspace dimension is known beforehand and the number of iterations is preset to be $M = \lceil 6/p \rceil$ $(p = 1, 2, 3)$.

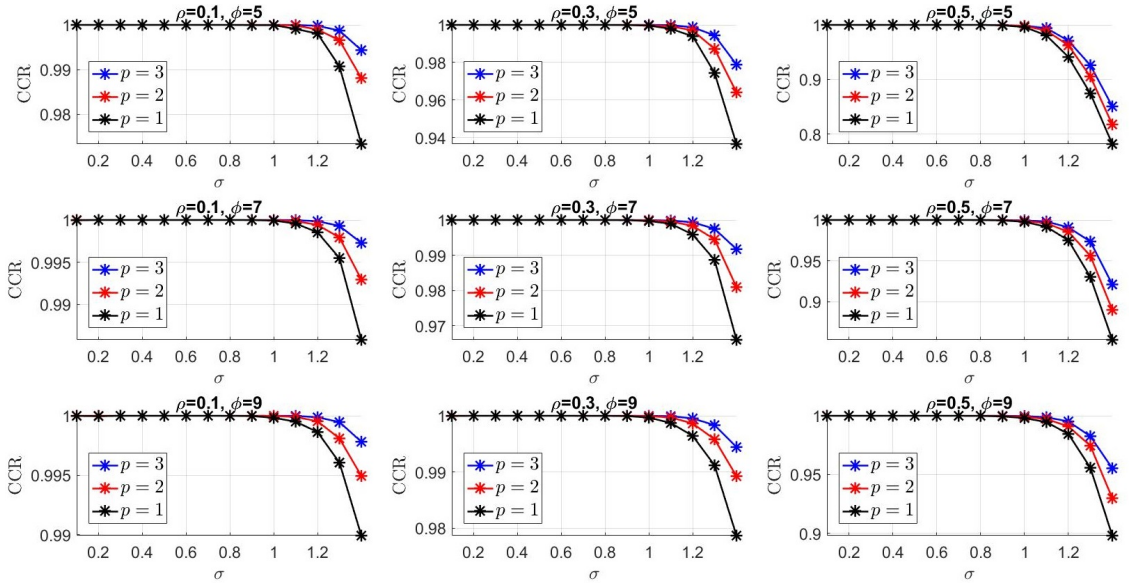

Figure 3: CCR versus noise standard deviation σ for nine different pairs of subspace affinity ρ and sample density ϕ; assume the subspace dimension is known beforehand and the number of iterations is preset to be $M = \lceil 6/p \rceil$ $(p = 1, 2, 3)$.

and $M = \lceil 6/p \rceil$ for GOMP, respectively; this represents the ideal situation that subspace dimension (equal to 6) is known. It is first seen from Fig. 2 that, as compared to the conventional OMP ($p = 1$), the proposed GOMP with multiple neighbor selection ($p > 1$) yields higher TNR. In particular, GOMP with $p = 3$ achieves the highest TNR because it involves fewest iterations (only two) and, thus, is subject to least perturbation of the residual vector. Then, it is seen from Fig. 3 that improved TNR in turn leads to higher CCR. The above experiment is conducted again by instead using the proposed stopping rule (5) to halt the algorithm, and the results are shown in Fig. 4 and Fig. 5. It can be seen that the proposed GOMP still outperforms

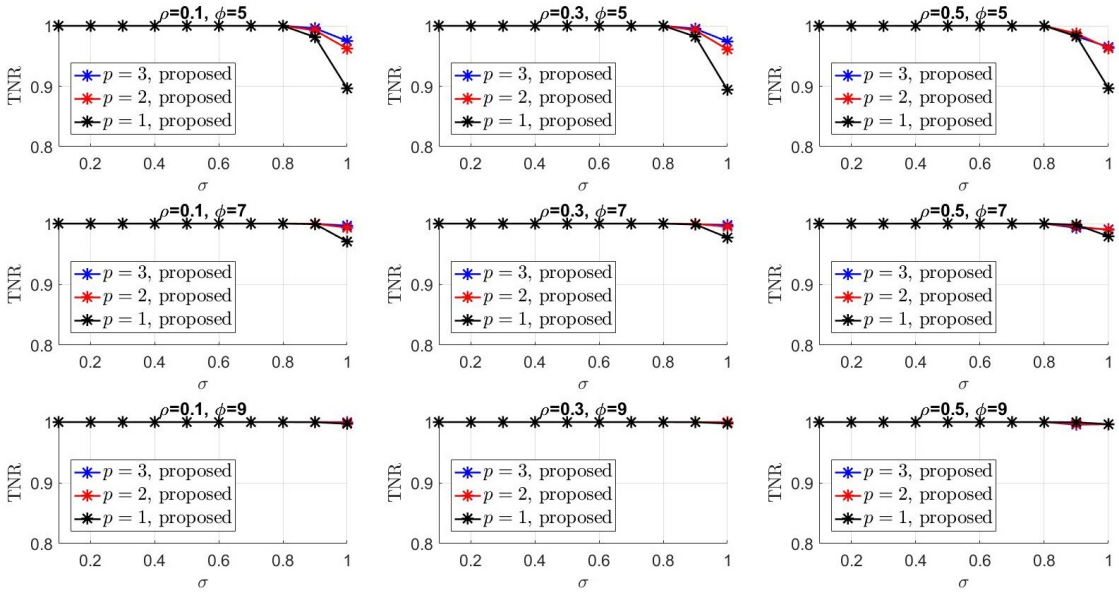

Figure 4: TNR versus noise standard deviation σ for nine different pairs of subspace affinity ρ and sample density ϕ; the proposed stopping rule (5) is used.

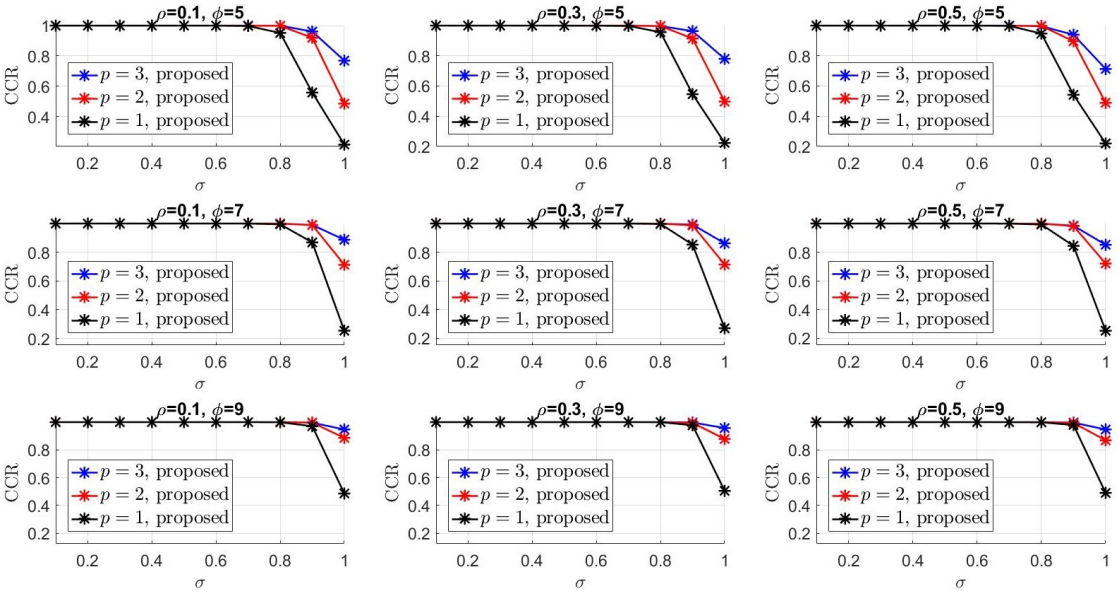

Figure 5: CCR versus noise standard deviation σ for nine different pairs of subspace affinity ρ and sample density ϕ; the proposed stopping rule (5) is used.

OMP. Fig. 6 further plots the ANRN of the two methods, both employing the proposed stopping rule (5), for different noise standard deviation σ. We first observe from the figure that, when noise is so small that $\sigma < 0.1$, ANRN of all cases is about six, in support of the assertion in Theorem 2 that as many neighbors as the subspace dimension suffice to well explain the data point under consideration. When σ grows up to 0.1, the proposed GOMP (with $p = 2, 3$) tends to retrieve more neighbors. As we have mentioned in the discussions at the end of Section 3, further noise corruption would enlarge the dimension of the span of the data cluster; as a result, the algorithm is terminated after more than $\lceil d_L/p \rceil$ iterations so as to recover more neighbors (than the ground truth subspace dimension) to explain the data point. When σ gets even larger,

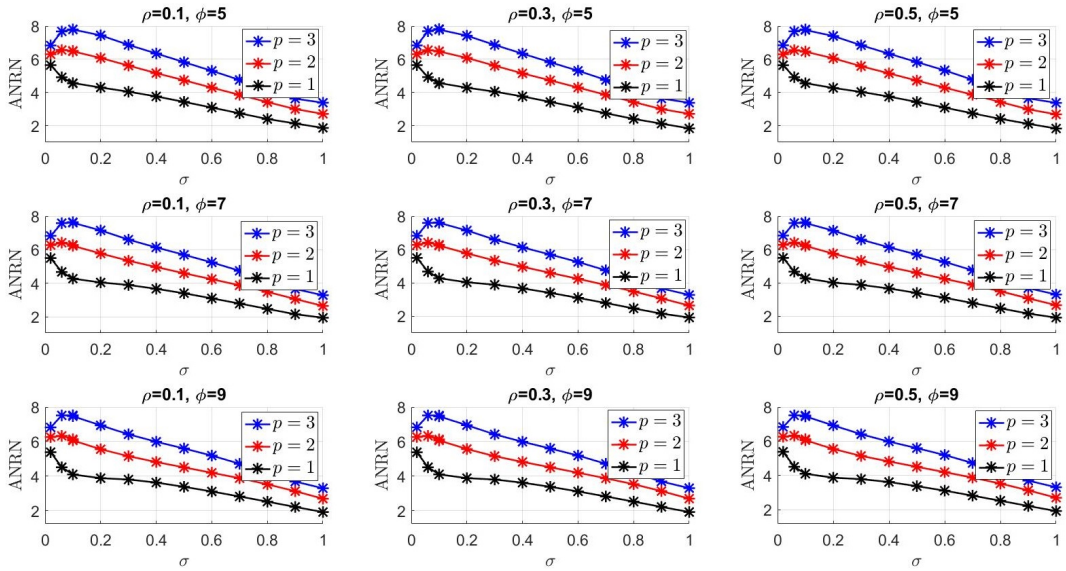

Figure 6: ANRN versus noise standard deviation σ for nine different pairs of subspace affinity ρ and sample density ϕ; the proposed stopping rule (5) is used.

Figure 7: Samples of four real datasets. (a) The Extended Yale B human face dataset. (b) The Coil-100 object image dataset. (c) The MNIST handwritten digit dataset. (d) The Devanagari handwritten character dataset.

the residual is then severely dominated by noise, only to halt neighbor search earlier, say, in less than $\lceil d_L/p \rceil$ iterations; this ends up with fewer recovered neighbors and reduced ANRN.

4.2 Real-World Data

We proceed to evaluate the performance of the proposed method by using the following four real-world datasets (an illustration of their test samples is shown in Fig. 7).

- The Extended Yale B human face dataset (Lee et al., 2005), comprised of photos of 38 people each with 65 images of 192×168 pixels. To reduce the computational cost, we use the dimensionality reduction technique as in (Elhamifar & Vidal, 2013; Peng et al., 2013; Matsushima & Brbic, 2019) for reducing the image size to 48×42.
- The Coil-100 dataset (Nene et al., 1996), composed of 7200 images (each with size 32×32) of 100 objects (72 images for each object).
- The MNIST dataset (LeCun et al., 1998), containing 70000 images of handwritten digits with size 28×28; at least 6000 images for each digit.
- The Devanagari dataset (Acharya et al., 2015), containing images of handwritten Devanagari characters and digits. We use the character image part, which consists of 72000 images of 36 characters with size 32×32 (2000 images per character).

Table 1: Running time (in minutes) of OMP ($p = 1$) and GOMP ($p = 2, 3, 4, 6$) when 12 neighbors are recovered.

Running time (min)	$p = 1$	$p = 2$	$p = 3$	$p = 4$	$p = 6$
Extended Yale B	1.51	0.85	0.61	0.48	0.3
Coil-100	2.83	1.45	0.98	0.75	0.51
MNIST	194	100	68	50	37
Devanagari	218	130	91	72	47

Table 2: TNR of OMP ($p = 1$) and GOMP ($p = 2, 3, 4, 6$) when 12 neighbors are recovered.

TNR	$p = 1$	$p = 2$	$p = 3$	$p = 4$	$p = 6$
Extended Yale B	0.37	0.41	0.46	0.49	0.58
Coil-100	0.20	0.21	0.26	0.33	0.45
MNIST	0.58	0.61	0.64	0.66	0.68
Devanagari	0.2	0.23	0.30	0.39	0.51

Table 3: FDR of OMP ($p = 1$) and GOMP ($p = 2, 3, 4, 6$) when 12 neighbors are recovered.

FDR	$p = 1$	$p = 2$	$p = 3$	$p = 4$	$p = 6$
Extended Yale B	0.85	0.86	0.87	0.87	0.90
Coil-100	0.94	0.93	0.93	0.94	0.95
MNIST	0.80	0.81	0.81	0.83	0.84
Devanagari	0.51	0.56	0.60	0.65	0.72

In the first part of simulations, we compare the proposed GOMP with OMP using the above four real datasets. Assuming that a total number of 12 neighbors are to be recovered, Table 1 lists the running time for OMP ($p = 1$) and GOMP ($p = 2, 3, 4, 6$); the experiments are conducted using MATLAB 2016b on a desktop with an AMD 3700X CPU of 3.6GHz and 64 GB RAM. It can be seen that GOMP enjoys shorter running time, confirming the computational efficiency of GOMP. Table 2 then compares the resultant TNR, showing that GOMP yields higher TNR when the same number of neighbors are to be found. We note that, unlike the synthetic data case, TNR of the two greedy methods are less than 0.5 for Yale-B, Coil-100, and Devanagari datasets, in disagreement with the high-TNR assertion of our recovery rate analyses in Section 3. This is because real-world datasets seldom meet the assumptions underlying the semi-random model (Assumptions 1 and 2): neither the signal points are uniform nor the noise is Gaussian (e.g., real human face data are typically corrupted by sparse outliers (Elhamifar & Vidal, 2013; Soltanolkotabi & Candès, 2012)). Hence, the developed recovery rate formulae in Section 3 may not fully explain the simulated results when real-world datasets are considered. Despite TNR may not be high in certain cases, we would like to remark that the existence of false neighbors does not largely degrade final data clustering accuracy. To see this, we further consider the feature detection rate (FDR) (Heckel & Bölcskei, 2015), defined to be

$$FDR \triangleq \frac{1}{N} \sum_{i=1}^{N} \frac{\|\widetilde{\mathbf{c}}_i^*\|_2}{\|\mathbf{c}_i^*\|_2}, \tag{22}$$

Table 4: CCR of OMP ($p = 1$) and GOMP ($p = 2, 3, 4, 6$) when 12 neighbors are recovered.

CCR	$p = 1$	$p = 2$	$p = 3$	$p = 4$	$p = 6$
Extended Yale B	0.47	0.49	0.56	0.63	0.67
Coil-100	0.29	0.32	0.37	0.42	0.48
MNIST	0.43	0.51	0.63	0.64	0.66
Devanagari	0.26	0.31	0.35	0.37	0.38

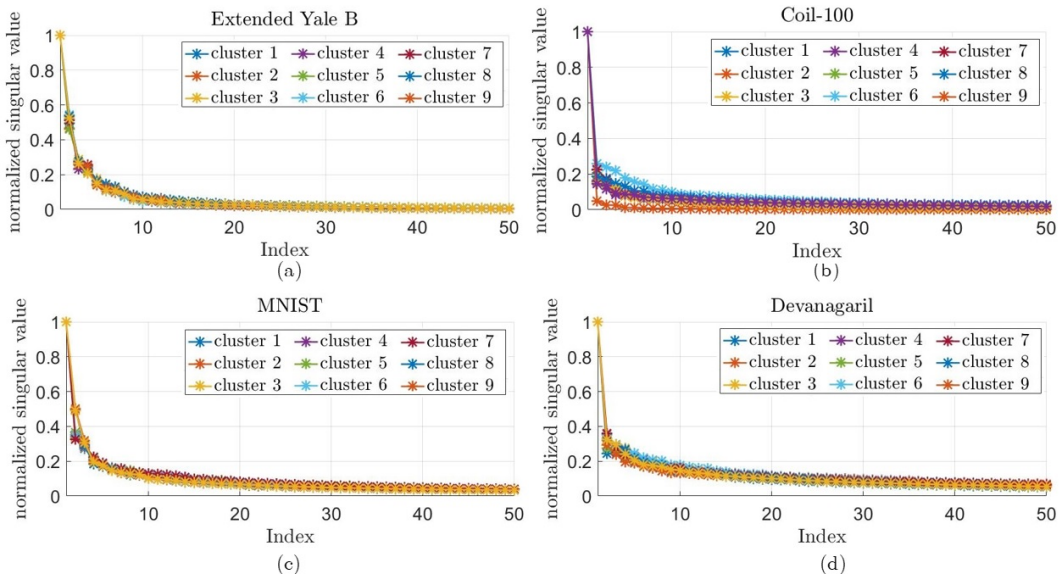

Figure 8: Normalized singular values of data matrices of the first 9 cluster of (a) Extended Yale B dataset, (b) Coil-100 dataset, (c) MNIST dataset, and (d) Devanagari dataset.

where $\widetilde{\mathbf{c}}_i^*$ is the vector containing the entries of $\mathbf{c}_i^*$ supported on true neighbors. If FDR is high, the weights on the false edges in the similarity graph are small, meaning that the existence of false neighbors is somehow not harmful. Table 3 presents the computed FDR; it can be observed that both OMP and GOMP achieve high FDR (above 0.7 for most cases). Table 4 then lists the CCR, showing that GOMP achieves higher CCR thanks to higher TNR.

Finally, we illustrate the ANRN of GOMP when employing the proposed stopping rule (5), and discuss the selection of the parameter p aimed at fulfilling the dimension-aware property. For this purpose, knowledge about ground truth subspace dimensions and noise level is needed. To determine the subspace dimensions, we vectorize all data points associated with each cluster, stack them into a matrix, and calculate the normalized singular values (with respective to the maximal one). For each dataset, we choose the first 9 clusters for singular value computation and the results are plotted in Fig. 8, respectively. Observe from the four figures that the normalized singular values of all datasets exhibit a sudden slump from 1 to 0.1, beyond which they decrease slowly; the curves are seen to bend to the floors around 7∼9 for Yale-B, 5∼7 for Coil-100, 10∼15 for MNIST, and 15∼20 for Devanagari, respectively. The above singular value index ranges are then used as the estimated subspace dimensions. Also, the noise power of each dataset is estimated as the averaged sum of the rest nondominant singular values. Table 5 then shows the ranges of estimated subspace dimensions, average subspace affinity, estimated noise powers, and the computed ANRN of GOMP using the proposed stopping rule (5) with $p = 2, 3, 4, 6$. We can first observe from the table that, for the Coil-100 and MNIST datasets (both with relatively small affinity, though the latter subject to large noise), their ANRN's are within the ranges 4.7∼12.1 and 8.6∼17.7, respectively, which are fairly close to the estimated subspace dimensions

Table 5: Estimated subspace dimensions, noise powers, subspace affinity, data size, and ANRN of GOMP with proposed stopping rule.

	Coil-100	MNIST	Devanagari	Extended Yale B
Range of subspace dimension	[6,10]	[10,15]	[15,20]	[7,9]
Estimated noise powers	0.036	0.194	0.199	0.032
Estimated subspace affinity	0.447	0.553	0.628	0.706
Data size	7200	70000	72000	2470
ANRN ($p = 2$)	4.7	8.6	6.2	7.5
ANRN ($p = 3$)	6.3	11.6	7.4	11.2
ANRN ($p = 4$)	8.4	13.5	9.4	14.3
ANRN ($p = 6$)	12.1	17.7	16.8	19.6

$[6, 10]$ and $[10, 15]$. This implies that, for these two datasets, the proposed stopping rule (5) is aware of the subspace dimension. However, for the Devanagari dataset (with a medium-to-large affinity) and Extended Yale B dataset (with a large affinity, despite small noise corruption), their ANRN's deviate largely from the subspace dimensions, and hence the dimension-aware property seems to fail. We particularly observe that, for the Devanagari dataset, the algorithm is halted earlier (in less than $\lceil d/p \rceil$ iterations) with a small p ($p = 2, 3, 4$), resulting in fewer recovered neighbors and reduced ANRN. This is mainly because, with large noise corruption, the residual is quickly dominated by noise so as to trigger the stopping rule (5) early (hence, less than $\lceil d/p \rceil$ iterations); in this case, a large p (say $p = 6$) to soon recover many neighbors is desired. Such a tendency is also seen in the synthetic data case (see Fig. 6). Regarding the Extended Yale B dataset, GOMP with a small p ($p = 2, 3$) is instead preferred. This is because, with a large affinity, different clusters are close to each other; if p is large, the algorithm is likely to pick neighbors from incorrect clusters and is then misled to find neighbors from some subspace (other than the ground truth one) with a larger dimension, thereby ending up with higher ANRN. Based on the above simulation findings we conclude that, as long as the subspace affinity is small, the proposed stopping rule (5) is aware of the subspace dimension, allowing for a wide range of p. If the affinity is medium-to-large, a good p is highly dependent on the noise power: large (small, respectively) p for large (small, respectively) noise power. Overall, selection of p is therefore closely related to the subspace affinity and noise power, which are nonetheless unknown unless additional off-line estimation is further conducted. Notably, we can also observe from Table 5 that the good p increases with the data size N; this has also been inferred based on our recovery rate analysis for the semi-random model (see the discussions in Section 3.2). Hence a very simple guideline, without the need of knowing the subspace dimensions and noise strength, is to choose small p (say, $p = 2, 3$) for small-size datasets (with thousands of data points or less) and large p ($p = 5, 6$) for large-scale datasets (with tens of thousands of data points or more).

In the second part of this simulation, we compare GOMP using the proposed stopping rule (5) (dubbed by "proposed") with the following neighbor identification schemes:

- GOMP terminated after $\lceil d/p \rceil$ iterations, in which d is the ground-truth subspace dimension determined as the floor of the average of the subspace dimensions within the range in Table 5. This serves as the ideal implementation of GOMP that employs the knowledge of the subspace dimension (hereafter coined as "ideal").

- OMP (Tschannen & Bölcskei, 2018) terminated after a pre-determined number M of iterations.

- Active OMP (AOMP) (Chen et al., 2018) terminated after a pre-determined number M of iterations; the neighbor dropping probability is set be 0.8, as used in (Chen et al., 2018).

Table 6: Simulation parameters for GOMP, OMP and AOMP.

Parameters	Extended Yale B	Coil-100	MNIST	Devanagari
GOMP-ideal (p)	2	3	4	6
GOMP-proposed (p)	2	3	4	6
OMP (M)	8	4	4	5
AOMP (M)	12	4	6	6

Table 7: CCR of different clustering algorithms (the blank denotes that time limit of 7 days is exceeded).

CCR	OMP	AOMP	ROMP	SSC	ESC	S3C	S5C	ORN	proposed	ideal
Yale B	0.73	0.81	0.77	0.81	0.44	0.33	0.71	0.83	0.84	0.84
Coil 100	0.46	0.35	0.38	0.79	0.23	0.14	0.53	0.76	0.71	0.79
MNIST	0.61	0.64	0.63	$\times$	0.27	0.22	0.55	0.72	0.64	0.66
Deva.	0.28	0.33	0.28	$\times$	0.20	0.15	0.33	0.40	0.35	0.38

- Restricted connection OMP (ROMP) (Zhu et al., 2019); following the same setting as in (Zhu et al., 2019), the number of iterations is set to be 3, and the number of restricted connections is set to be 2.
- SSC with conventional ℓ_1-minimization (LASSO) technique (Elhamifar & Vidal, 2013); the regularization parameter λ is obtained via an exhaustive search over the interval $(0, 0.5]$, as in (Matsushima & Brbic, 2019).
- Exemplar-based subspace clustering (You et al., 2018); the number of exemplar data points is set to be 160, as used in (You et al., 2018), and the regularization parameter λ is obtained via an exhaustive search over the interval $(0, 0.5]$, as in (Matsushima & Brbic, 2019).
- Scalable sparse subspace clustering (S3C) (Peng et al., 2013); following the same setting as in (Peng et al., 2013), we randomly choose 1000 data points as "in-sample data" to find coarse estimates of the ground-truth subspaces; the regularization parameter λ is obtained via an exhaustive search over $(0, 0.5]$, as in (Matsushima & Brbic, 2019).
- Selective sampling-based scalable sparse subspace clustering (S5C) (Matsushima & Brbic, 2019); to carry out representation learning to find representative data points, the batch size is set to be 1, as used in (Matsushima & Brbic, 2019), and the size of representation set is chosen to be $10L$. Again, the regularization parameter λ is obtained via an exhaustive search over $(0, 0.5]$ as in (Matsushima & Brbic, 2019).
- Elastic net subspace clustering with oracle guided elastic net (ORN) (You et al., 2016a); the noise trade off parameter γ is set to be 1, and the regularization parameter λ is obtained via an exhaustive search over $(0, 0.5]$, as in (Matsushima & Brbic, 2019).

Note that the last four methods, namely, ESC, S3C, S5C, and ORN, are reduced-complexity variants of the ℓ_1-minimization scheme (Elhamifar & Vidal, 2013). For both OMP and AOMP, we conduct an exhaustive search over the integer set $\{1, 2, ..., 18\}$ to find the best number M of iterations that achieve highest CCR; such an approach has been adopted in the simulation study of (Matsushima & Brbic, 2019); for both GOMP-proposed and GOMP-ideal, the parameter p for each dataset is selected according to the suggested guidelines (see Table 6 for a list of the parameters used for OMP and GOMP algorithms). Table 7 shows the CCR of all methods. Compared to the competing greedy algorithms OMP, AOMP, and ROMP, GOMP-proposed

Table 8: Running time (in minutes) of different clustering algorithms (the blank denotes that time limit of 7 days is exceeded).

(minute)	OMP	AOMP	ROMP	SSC	ESC	S3C	S5C	ORN	proposed	ideal
Yale B	0.96	1.55	0.87	577	4	140	11	13	0.55	0.43
Coil 100	0.95	2.7	0.82	988	6	138	42	31	0.55	0.5
MNIST	77	56	54	×	83	109	46	69	52	37
Deva.	96	74	69	×	96	142	54	80	59	47

(proposed) achieves higher CCR. As against the five ℓ_1-minimization based solutions, GOMP-proposed (proposed) is next to ORN (You et al., 2016a), which is essentially a LASSO regression modified to further promote connectivity. It is also observed that GOMP-proposed (proposed) performs just slightly inferior to GOMP-ideal (ideal), which requires knowledge of the subspace dimension. This indicates that the proposed stopping rule (5) in conjunction with the suggested p is aware of the subspace dimension. Table 8 then lists the running time of all methods. The results show that GOMP-proposed (proposed) is very computationally efficient: it enjoys the lowest running time in most cases. For the MNIST and Devanagari datasets, S5C is faster, but is inferior to GOMP-ideal (ideal).

5 Proofs

5.1 Proof of Theorem 1

We set about the proof by defining the following per-iteration neighbor recovery event:

$$E_m \triangleq \{\text{at least } k_m \text{ true neighbors are obtained in the } m\text{th iteration}\},\ 1 \leq m \leq M. \tag{23}$$

It suffices to show the event $\bigcap_{m=1}^{M} E_m$ holds with a probability as high as claimed in (11). The following lemmas are needed for deriving Theorem 1.

Lemma 1. *((Davenport et al., 2011, Ex. 25); (Heckel & Bölcskei, 2015, eq. 59)) Let $\mathbf{a} \in \mathbb{R}^m$ be uniformly distributed over the unit-sphere of $\mathbb{R}^m$, and $\mathbf{b} \in \mathbb{R}^m$ be a random vector independent of $\mathbf{a}$. Then, for $\epsilon \geq 0$ we have*

$$Pr\{|\mathbf{a}^T\mathbf{b}| > \epsilon\|\mathbf{b}\|_2|\} \leq 2e^{-m\epsilon^2/2}, \tag{24}$$

$$Pr\{|\mathbf{a}^T\mathbf{b}| < \epsilon\|\mathbf{b}\|_2|\} \leq \sqrt{2/\pi}\epsilon. \tag{25}$$

Lemma 2. *Let $\mathbf{U}_l \in \mathbb{R}^{n\times d_l}$ be a matrix whose columns form an orthonormal basis for subspace $\mathcal{S}_l$, $1 \leq l \leq L$. For $\mathbf{x}_i \notin \mathcal{S}_L$, we have*

$$Pr\left\{\left|\left\langle \mathbf{x}_i, \frac{\mathbf{r}_{m,\|}^{(N)}}{\|\mathbf{r}_{m,\|}^{(N)}\|_2}\right\rangle\right| > \max_{l\neq L} 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l d_L}}\right\} < \frac{2}{N^{(8\log N)/d_L}} \tag{26}$$

Proof : See Appendix A. □

Lemma 3. *For $\mathbf{x}_i \notin \mathcal{S}_L$, we have*

$$Pr\left\{\left|\left\langle \mathbf{x}_i, \frac{\mathbf{r}_{m,\perp}^{(N)}}{\|\mathbf{r}_{m,\perp}^{(N)}\|_2}\right\rangle\right| \leq \sqrt{\frac{6\log N}{n-d_L}}\right\} \geq 1 - 2cN^{-3} \tag{27}$$

Proof : See Appendix B. □

If $k_m = 0$, it follows $\Pr\{E_m\} = 1$ because the number of true neighbors is never negative. We then consider the case that $k_m > 0$. Since the already chosen data vectors in all the previous $m-1$ iterations are orthogonal

to the residual $\mathbf{r}_{m-1}^{(N)}$, i.e., $|\langle \mathbf{y}_j, \mathbf{r}_{m-1}^{(N)}\rangle| = 0$ for all $j \in \Lambda_{m-1}$, the "yet-to-be-selected" candidate neighbors are $\mathbf{y}_j$'s for $j \in \{1, 2, ..., N-1\}\backslash\Lambda_{m-1}$. Given that $\Lambda_{m-1} = \Lambda_{m-1}^t \cup \Lambda_{m-1}^f$, where the disjoint subsets Λ_{m-1}^t and Λ_{m-1}^f consist of, respectively, the indexes of true and false neighbors already selected up to the first $m-1$ iterations, $|\mathcal{Y}_L| - 1 - |\Lambda_{m-1}^t|$ and $N - |\mathcal{Y}_L| - |\Lambda_{m-1}^f|$ data vectors remain in the same and different clusters as $\mathbf{y}_N$, respectively. Let the true neighbor candidates $\mathbf{y}_{t_j}$, where $\mathbf{x}_{t_j} \in \mathcal{S}_L$ and $t_j \notin \Lambda_{m-1}^t$, be sorted according to $|\langle \mathbf{y}_{t_1}, \mathbf{r}_{m-1}^{(N)}\rangle| \geq ... \geq |\langle \mathbf{y}_{t_{|\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|}}, \mathbf{r}_{m-1}^{(N)}\rangle|$, and likewise the false neighbor candidates $\mathbf{y}_{f_k}$, where $\mathbf{x}_{f_k} \notin \mathcal{S}_L$ and $f_k \notin \Lambda_{m-1}^f$, in the way that $|\langle \mathbf{y}_{f_1}, \mathbf{r}_{m-1}^{(N)}\rangle| \geq ... \geq |\langle \mathbf{y}_{f_{N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|}}, \mathbf{r}_{m-1}^{(N)}\rangle|$. Also, let the noiseless signal vectors $\mathbf{x}_{\bar{t}_g} \in \mathcal{S}_L$, $\bar{t}_g \notin \Lambda_{m-1}^t$, be ordered so that $|\langle \mathbf{x}_{\bar{t}_1}, \mathbf{r}_{m-1}^{(N)}\rangle| \geq ... \geq |\langle \mathbf{x}_{\bar{t}_{|\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|}}, \mathbf{r}_{m-1}^{(N)}\rangle|$, while those $\mathbf{x}_{\bar{f}_h} \notin \mathcal{S}_L$, $\bar{f}_h \notin \Lambda_{m-1}^f$, in the way $|\langle \mathbf{x}_{\bar{f}_1}, \mathbf{r}_{m-1}^{(N)}\rangle| \geq ... \geq |\langle \mathbf{x}_{\bar{f}_{N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|}}, \mathbf{r}_{m-1}^{(N)}\rangle|$.

The event E_m defined in (23) can be expressed in terms of the above ordered statistics. By definition, E_m occurs when at least k_m true neighbors, and so at most $p - k_m$ false neighbors, are recovered in the mth iteration. That is to say, $\mathbf{y}_{t_{k_m}}$ is chosen as a neighbor but $\mathbf{y}_{f_{p-k_m+1}}$ is not, justifying the inequality $|\langle \mathbf{y}_{f_{p-k_m+1}}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle|$. The converse is obviously true. Hence we can rewrite E_m as

$$E_m = \{|\langle \mathbf{y}_{f_{p-k_m+1}}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle|\}. \tag{28}$$

The expression (28) involves ordered statistics of the absolute inner products between the residual and noisy data points $\mathbf{y}_j$'s, whose distribution is however quite complicated. To ease our lower bound derivation, we shall instead seek sufficient conditions, specified by absolute inner products between the residual and noise-free signal points $\mathbf{x}_j$'s, for $|\langle \mathbf{y}_{f_{p-k_m+1}}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle|$; this allows us to employ the assumed uniform distribution of $\mathbf{x}_j$'s (Assumption 1) as well as the ground-truth subspace orientation (Assumption 3) to facilitate analysis. For this we first note from (8) that, for all $1 \leq j \leq N-1$, we have

$$\langle \mathbf{y}_j, \mathbf{r}_{m-1}^{(N)}\rangle = \langle \mathbf{x}_j, \mathbf{r}_{m-1,\|}^{(N)}\rangle + \langle \mathbf{x}_j, \mathbf{r}_{m-1,\perp}^{(N)}\rangle + \langle \mathbf{e}_j, \mathbf{r}_{m-1}^{(N)}\rangle. \tag{29}$$

Then for all $u \leq k_m$, it follows

$$\begin{aligned} |\langle \mathbf{y}_{\bar{t}_u}, \mathbf{r}_{m-1}^{(N)}\rangle| &= |\langle \mathbf{x}_{\bar{t}_u}, \mathbf{r}_{m-1,\|}^{(N)}\rangle + \langle \mathbf{x}_{\bar{t}_u}, \mathbf{r}_{m-1,\perp}^{(N)}\rangle + \langle \mathbf{e}_{\bar{t}_u}, \mathbf{r}_{m-1}^{(N)}\rangle| \\ &\overset{(a)}{=} |\langle \mathbf{x}_{\bar{t}_u}, \mathbf{r}_{m-1,\|}^{(N)}\rangle + \langle \mathbf{e}_{\bar{t}_u}, \mathbf{r}_{m-1}^{(N)}\rangle| \\ &\geq |\langle \mathbf{x}_{\bar{t}_u}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| - \max_{1\leq j\leq |\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|} |\langle \mathbf{e}_{t_j}, \mathbf{r}_{m-1}^{(N)}\rangle| \\ &\overset{(b)}{\geq} |\langle \mathbf{x}_{\bar{t}_{k_m}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| - \max_{1\leq j\leq |\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|} |\langle \mathbf{e}_{t_j}, \mathbf{r}_{m-1}^{(N)}\rangle|, \end{aligned} \tag{30}$$

where (a) holds since $\mathbf{x}_{t_j} \in \mathcal{S}_L$ for $\mathbf{x}_{t_j} \in \mathcal{S}_L$, and (b) follows from the ordering of $\mathbf{x}_{\bar{t}_g}$'s. Hence, at least k_m absolute inner products $|\langle \mathbf{y}_{\bar{t}_u}, \mathbf{r}_{m-1}^{(N)}\rangle|$ are no smaller than the term on the right-hand-side (RHS) of (30); in particular, for $|\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle|$, the k_mth largest $|\langle \mathbf{y}_{t_j}, \mathbf{r}_{m-1}^{(N)}\rangle|$, we must have

$$|\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle| \geq |\langle \mathbf{x}_{\bar{t}_{k_m}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| - \max_{1\leq j\leq |\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|} |\langle \mathbf{e}_{t_j}, \mathbf{r}_{m-1}^{(N)}\rangle|. \tag{31}$$

Similarly, by the ordering of $|\langle \mathbf{x}_{\bar{f}_k}, \mathbf{r}_{m-1,\|}^{(N)}\rangle|$, for all $u \geq p - k_m + 1$, we have

$$\begin{aligned} |\langle \mathbf{y}_{f_{p-k_m+1}}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{x}_{\bar{f}_{p-k_m+1}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| &+ \max_{1\leq k\leq N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|} |\langle \mathbf{x}_{f_k}, \mathbf{r}_{m-1,\perp}^{(N)}\rangle| \\ &+ \max_{1\leq k\leq N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|} |\langle \mathbf{e}_{f_k}, \mathbf{r}_{m-1}^{(N)}\rangle|. \end{aligned} \tag{32}$$

Putting (31) and (32) together, the condition $|\langle \mathbf{y}_{f_{p-k_m+1}}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{y}_{t_{k_m}}, \mathbf{r}_{m-1}^{(N)}\rangle|$ is guaranteed, and hence the event E_m occurs, once the following inequality is true

$$\begin{aligned} &|\langle \mathbf{x}_{\overline{f}_{p-k_m+1}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| + \max_{1\leq k\leq N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|} |\langle \mathbf{x}_{f_k}, \mathbf{r}_{m-1,\perp}^{(N)}\rangle| \\ &+ \max_{1\leq k\leq N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|} |\langle \mathbf{e}_{f_k}, \mathbf{r}_{m-1}^{(N)}\rangle| + \max_{1\leq j\leq |\mathcal{Y}_L|-1-|\Lambda_{m-1}^t|} |\langle \mathbf{e}_{t_j}, \mathbf{r}_{m-1}^{(N)}\rangle| \leq |\langle \mathbf{x}_{\overline{t}_{k_m}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle|. \end{aligned} \tag{33}$$

Recall that our goal is to show that $\bigcap_{m=1}^M E_m$ occurs with a high probability. Under the semi-random model, we go on to estimate the probability of which the inequality (33) holds for all $1 \leq m \leq M$, and in turn a lower bound for $\Pr\{\bigcap_{m=1}^M E_m\}$. The basic idea is to "split and then lump": find an upper bound for each individual left-hand-side (LHS) term of (33) and estimate, one by one, the probability about which the obtained inequality holds, and similarly a lower bound for the RHS term along with an estimated probability of its validity; putting the results altogether gives a sufficient condition for (33) along with the desired probability lower bound.

a) An Upper Bound for the 1st LHS Term of (33): For a fixed α, we first note that

$$\begin{aligned} &\Pr\{|\langle \mathbf{x}_{\overline{f}_{p-k_m+1}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| \leq \alpha\} \\ &= 1 - \Pr\{|\langle \mathbf{x}_{\overline{f}_{p-k_m+1}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| > \alpha\} \\ &\overset{(a)}{\geq} 1 - \Pr\big\{\exists I \subset [N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|] \text{ with} \\ &\qquad\qquad |I| = p-k_m+1 \text{ s.t. } |\langle \mathbf{x}_{f_k}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| > \alpha, \forall k \in I\big\} \\ &\overset{(b)}{\geq} 1 - \binom{N-|\mathcal{Y}_L|}{p-k_m+1}\Big(\max_{\substack{I:|I|=p-k_m+1 \\ I\subset[N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|]}} \Pr\{|\langle \mathbf{x}_{f_k}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| > \alpha, \forall k \in I\}\Big) \\ &\overset{(c)}{\geq} 1 - \left(\frac{e(N-|\mathcal{Y}_L|)}{p-k_m+1}\right)^{p-k_m+1} \Big(\max_{\substack{I:|I|=p-k_m+1 \\ I\subset[N-|\mathcal{Y}_L|-|\Lambda_{m-1}^f|]}} \Pr\{|\langle \mathbf{x}_{f_k}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| > \alpha, \forall k \in I\}\Big), \end{aligned} \tag{34}$$

where in (a) the notation $[Q] \triangleq \{1, 2, ..., Q\}, Q \in \mathbb{N}$, in (b) $\binom{Q}{k} \triangleq \frac{Q!}{(Q-k)!k!}$, and (c) holds by (Cormen et al., 2001). By Lemma 1 and since data points are independent, for $|I| = p - k_m + 1$ we have

$$\Pr\left\{|\langle \mathbf{x}_{f_k}, \mathbf{r}_{m,\|}^{(N)}\rangle| > \max_{l\neq L} 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l d_L}}\|\mathbf{r}_{m,\|}^{(N)}\|_2, \forall k \in I\right\} < (\frac{2}{N^{(8\log N)/d_L}})^{p-k_m+1}. \tag{35}$$

Combining (34) and (35) and setting $\alpha = \max_{l\neq L} 4\log(N)\|\mathbf{U}_l^T\mathbf{U}_L\|_F\|\mathbf{r}_{m,\|}^{(N)}\|_2 \,/\sqrt{d_l d_L}$, it then follows

$$|\langle \mathbf{x}_{\overline{f}_{p-k_m+1}}, \mathbf{r}_{m-1,\|}^{(N)}\rangle| \leq \max_{l\neq L} 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l d_L}}\|\mathbf{r}_{m,\|}^{(N)}\|_2 \tag{36}$$

holds with a probability at least $1 - (\frac{2e(N-|\mathcal{Y}_L|)}{(p-k_m+1)N^{(8\log N)/d_L}})^{p-k_m+1}$.

b) An Upper Bound for the 2nd LHS Term of (33): We first note that

$$\begin{aligned}
&\Pr\Big\{\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|\le\alpha\Big\}\\
&=1-\Pr\Big\{\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|>\alpha\Big\}\\
&\ge 1-\Pr\{\exists k\in[N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|]\text{ s.t. }|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|>\alpha\}\\
&=1-\Pr\left\{\left(\bigcap_{k=1}^{N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\{|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|\le\alpha\}\right)^c\right\}\\
&=1-\Pr\left\{\bigcup_{k=1}^{N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\{|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|>\alpha\}\right\}\\
&\ge 1-\sum_{k=1}^{N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\Pr\{|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|>\alpha\}.
\end{aligned}\tag{37}$$

By Lemma 2, we have

$$\Pr\left\{|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m,\perp}\rangle|>\sqrt{\frac{6\log N}{n-d_L}}\|\mathbf{r}^{(N)}_{m,\perp}\|_2\right\}<2cN^{-3},\ 1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|.\tag{38}$$

Combing (37) and (38) and with $\alpha=\sqrt{\frac{6\log N}{n-d_L}}\|\mathbf{r}^{(N)}_{m,\perp}\|_2$, we obtain

$$\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}|\langle \mathbf{x}_{f_k},\mathbf{r}^{(N)}_{m-1,\perp}\rangle|\le\sqrt{\frac{6\log N}{n-d_L}}\|\mathbf{r}^{(N)}_{m,\perp}\|_2\tag{39}$$

holds with a probability at least $1-2cN^{-2}$.

c) An Upper Bound for the 3rd LHS Term of (33): Using similar techniques as in deriving (37) we can first reach

$$\Pr\Big\{\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}|\langle \mathbf{e}_{f_k},\mathbf{r}^{(N)}_{m-1}\rangle|\le\alpha\Big\}\ge 1-\sum_{k=1}^{N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\Pr\{|\langle \mathbf{e}_{f_k},\mathbf{r}^{(N)}_{m-1}\rangle|>\alpha\}.\tag{40}$$

Using (24) and with $\mathbf{a}=\mathbf{e}_{f_k}/\|\mathbf{e}_{f_k}\|_2$ and $\mathbf{b}=\mathbf{r}^{(N)}_{m-1}$, it follows

$$\begin{aligned}
&\Pr\left\{|\langle \mathbf{e}_{f_k},\mathbf{r}^{(N)}_{m}\rangle|>\sqrt{\frac{6\log N}{n}}\|\mathbf{r}^{(N)}_{m}\|_2\|\mathbf{e}_{f_k}\|_2\right\}\\
&=\Pr\left\{\left|\left\langle\frac{\mathbf{e}_{f_k}}{\|\mathbf{e}_{f_k}\|_2},\mathbf{r}^{(N)}_{m}\right\rangle\right|>\sqrt{\frac{6\log N}{n}}\|\mathbf{r}^{(N)}_{m}\|_2\right\}\le 2N^{-3}.
\end{aligned}\tag{41}$$

Under Assumption 1 we have

$$\|\mathbf{r}^{(N)}_m\|_2=\|\mathrm{P}_{\mathcal{R}(\mathbf{Y}_{\Lambda_{m-1}})}\mathbf{y}_N\|_2\le\|\mathbf{y}_N\|_2=\|\mathbf{x}_N+\mathbf{e}_N\|_2\le 1+\|\mathbf{e}_N\|_2,\tag{42}$$

in which $\mathcal{R}(\mathbf{Y}_{\Lambda_{m-1}})$ is the column space of $\mathbf{Y}_{\Lambda_{m-1}}$ and the last inequality holds by the triangle inequality. Combing (41) and (42) yields

$$\Pr\left\{|\langle \mathbf{e}_{f_k},\mathbf{r}^{(N)}_{m}\rangle|>\sqrt{\frac{6\log N}{n}}(1+\|\mathbf{e}_N\|_2)\|\mathbf{e}_{f_k}\|_2\right\}\le 2N^{-3}.\tag{43}$$

Setting $\alpha = \sqrt{\frac{6\log N}{n}}(1+\|\mathbf{e}_N\|_2)\max\limits_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\|\mathbf{e}_{f_k}\|_2$, (40) and (43) imply

$$\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}|\langle \mathbf{e}_{f_k}, \mathbf{r}^{(N)}_{m-1}\rangle| \le \sqrt{\frac{6\log N}{n}}(1+\|\mathbf{e}_N\|_2)\max_{1\le k\le N-|\mathcal{Y}_L|-|\Lambda^f_{m-1}|}\|\mathbf{e}_{f_k}\|_2 \tag{44}$$

holds with a probability at least $1-2N^{-2}$.

d) An Upper Bound for the 4th LHS Term of (33): By following the same procedures as from (40) to (44), we can readily show the inequality

$$\max_{1\le j\le|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|}|\langle \mathbf{e}_{t_j}, \mathbf{r}^{(N)}_{m-1}\rangle| \le \sqrt{\frac{6\log N}{n}}(1+\|\mathbf{e}_N\|_2)\max_{1\le j\le|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|}|\mathbf{e}_{t_j}\|_2. \tag{45}$$

e) A Lower Bound for the RHS Term of (33): Next, we go on to find a lower bound for the RHS term of (33). Using (25) and with $\mathbf{a}=\mathbf{x}_{t_j}$ and $\mathbf{b}=\mathbf{r}^{(N)}_{m-1,\|}$, we have

$$\Pr\left\{|\langle \mathbf{x}_{t_j}, \mathbf{r}^{(N)}_{m,\|}\rangle| < \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2\right\} \le \sqrt{\frac{2}{\pi}}\tau,\ 1\le j\le|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|, \tag{46}$$

which together with the assumption that $\mathbf{x}_{t_j}$'s are independent yields

$$\begin{aligned}\Pr\Big\{|\langle \mathbf{x}_{t_j}, \mathbf{r}^{(N)}_{m,\|}\rangle| < \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2,\ \forall j\in I\Big\} &\le \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m},\\ I\subset\big[|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|\big],\ |I| &= |\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m.\end{aligned} \tag{47}$$

It then follows

$$\begin{aligned}&\Pr\Big\{\langle \mathbf{x}_{t_{k_m}}, \mathbf{r}^{(N)}_{m-1,\|}\rangle| \ge \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2\Big\}\\
&= 1-\Pr\Big\{\langle \mathbf{x}_{t_{k_m}}, \mathbf{r}^{(N)}_{m-1,\|}\rangle| < \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2\Big\}\\
&\ge 1-\Pr\Big\{\exists I\subset\big[|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|\big] \text{ with } |I|=|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|-k_m+1\\
&\qquad\qquad \text{s.t. } |\langle \mathbf{x}_{t_j}, \mathbf{r}^{(N)}_{m,\|}\rangle| < \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2,\ \forall j\in I\Big\}\\
&\ge 1-\binom{|\mathcal{Y}_L|-1}{k_m-1}\max_{\substack{|I|=|\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m\\ I\subset[|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|]}}\Pr\Big\{|\langle \mathbf{x}_{t_j}, \mathbf{r}^{(N)}_{m,\|}\rangle| < \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m,\|}\|_2,\ \forall j\in I\Big\}\\
&\overset{(a)}{\ge} 1-\binom{|\mathcal{Y}_L|-1}{k_m-1}\max_{\substack{|I|=|\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m\\ I\subset[|\mathcal{Y}_L|-1-|\Lambda^t_{m-1}|]}}\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m}\\
&\ge 1-\left(\frac{e(|\mathcal{Y}_L|-1)}{k_m-1}\right)^{k_m-1}\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-k_m},\end{aligned} \tag{48}$$

where (a) holds thanks to (47). Hence, the inequality

$$|\langle \mathbf{x}_{\bar{t}_{k_m}}, \mathbf{r}^{(N)}_{m-1,\|}\rangle| \ge \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}^{(N)}_{m-1,\|}\|_2 \tag{49}$$

holds with a probability at least $1-\left(\frac{e(|\mathcal{Y}_L|-1)}{k_m-1}\right)^{k_m-1}\left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-|\Lambda^t_{m-1}|-k_m}$.

With the bounds in (36), (39), (44), (45) and (49), the inequality in (33) is true, and hence the event E_m occurs, when the following inequality holds:

$$\max_{l\neq L} 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l d_L}}\|\mathbf{r}_{m,\|}^{(N)}\|_2 + \sqrt{\frac{6\log N}{n-d_L}}\|\mathbf{r}_{m,\perp}^{(N)}\|_2 + 2\sqrt{\frac{6\log N}{n}}(1+\|\mathbf{e}_N\|_2)\max_{1\leq i\leq N}\|\mathbf{e}_i\|_2 \leq \frac{\tau}{\sqrt{d_L}}\|\mathbf{r}_{m-1,\|}^{(N)}\|_2, \tag{50}$$

or equivalently,

$$\max_{l\neq L}\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l}} + \frac{\sqrt{3d_L}\|\mathbf{r}_{m,\perp}^{(N)}\|_2}{\sqrt{8(n-d_L)\log N}\|\mathbf{r}_{m,\|}^{(N)}\|_2} + \frac{2\sqrt{3d_L}(1+\|\mathbf{e}_N\|_2)}{\sqrt{8n\log N}\|\mathbf{r}_{m-1,\|}^{(N)}\|_2}\max_{1\leq i\leq N}\|\mathbf{e}_i\|_2 \leq \frac{\tau}{4\log N}, \tag{51}$$

which is obtained by multiplying the inequality (50) throughout by the factor $\sqrt{d_L}\times 1/(4\log(N)\|\mathbf{r}_{m-1,\|}^{(N)}\|_2)$. It can then be concluded that occurs once (36), (39), (44), (45), (49) and (51) hold for all $1\leq m\leq M$. Hence, a lower bound for $\Pr\{\bigcap_{m=1}^M E_m\}$ can be obtained by finding a lower bound for the probability that (51) holds for all $1\leq m\leq M$.

Still, our approach seeks an upper bound for the LHS term in (51), hence a sufficient condition guaranteeing (51), and proves this bound holds with a high probability. Towards this end, we derive an upper bound for $\|\mathbf{r}_{m-1,\perp}^{(N)}\|_2$, $1\leq m\leq M$, an upper bound for $\|\mathbf{e}_i^{(N)}\|_2$, $1\leq i\leq N$, and a lower bound for $\|\mathbf{r}_{m-1,\|}^{(N)}\|_2$, $1\leq m\leq M$; lumping these altogether yields the claimed result. To begin with, we write

$$\|\mathbf{r}_{m-1,\perp}^{(N)}\|_2 = \|\mathrm{P}_{\mathcal{S}_L^\perp}(\mathbf{I}-\mathbf{Y}_{\Lambda_{m-1}}\mathbf{Y}_{\Lambda_{m-1}}^\dagger)\mathbf{y}_N\|_2 \leq \|\mathrm{P}_{\mathcal{S}_L^\perp}\mathbf{y}_N\|_2 \leq \|\mathbf{e}_N\|_2. \tag{52}$$

According to (Tschannen & Bölcskei, 2018, Lemma 10), the event $\bigcap_{i=1}^N\{\|\mathbf{e}_i\|_2\leq 3\sigma/2\}$ occurs with a probability at least $1-Ne^{-n/8}$. Hence, we conclude that the following sets of inequalities

$$\|\mathbf{r}_{m-1,\perp}^{(N)}\|_2 \leq 3\sigma/2, 1\leq m\leq M, \tag{53}$$

$$\|\mathbf{e}_i\|_2 \leq 3\sigma/2, 1\leq i\leq N, \tag{54}$$

hold at once with a probability at least $1-Ne^{-n/8}$. Next, we will find a lower bound for $\|\mathbf{r}_{m-1,\|}^{(N)}\|_2$'s. For $1\leq m\leq M$, we have

$$\begin{aligned}\mathbf{r}_{m-1,\|}^{(N)} &= \mathrm{P}_{\mathcal{S}_L}(\mathbf{I}-\mathbf{Y}_{\Lambda_{m-1}}\mathbf{Y}_{\Lambda_{m-1}}^\dagger)\mathbf{y}_N \\ &= \mathrm{P}_{\mathcal{S}_L}(\mathbf{y}_N-\mathbf{Y}_{\Lambda_{m-1}}\mathbf{c}_{m-1}^*) = \mathbf{y}_{N,\|}-\mathbf{Y}_{\Lambda_{m-1,\|}}\mathbf{c}_{m-1}^*,\end{aligned} \tag{55}$$

where $\mathbf{y}_{N,\|}\triangleq \mathrm{P}_{\mathcal{S}_L}\mathbf{y}_N$, $\mathbf{Y}_{\Lambda_{m-1,\|}}\triangleq \mathrm{P}_{\mathcal{S}_L}\mathbf{Y}_{\Lambda_{m-1}}$ and $\mathbf{c}_{m-1}^*\triangleq \underset{\mathbf{c}_{m-1}}{\operatorname{argmin}}\|\mathbf{y}_N-\mathbf{Y}_{\Lambda_{m-1}}\mathbf{c}_{m-1}\|_2$. Then a lower bound for $\|\mathbf{r}_{m-1,\|}^{(N)}\|_2$ is obtained as

$$\begin{aligned}\|\mathbf{r}_{m-1,\|}^{(N)}\|_2 &= \|\mathbf{y}_{N,\|}-\mathbf{Y}_{\Lambda_{m-1,\|}}\mathbf{c}_{m-1}^*\|_2 \\ &\geq \min_{\mathbf{c}}\|\mathbf{y}_{N,\|}-\mathbf{Y}_{\Lambda_{m-1,\|}}\mathbf{c}\|_2 \\ &= \|\mathbf{y}_{N,\|}-\mathbf{Y}_{\Lambda_{m-1,\|}}\mathbf{Y}_{\Lambda_{m-1,\|}}^\dagger\mathbf{y}_{N,\|}\|_2 \\ &\geq \|\mathbf{y}_{N,\|}-\mathbf{Y}_{\Lambda_{M-1,\|}}\mathbf{Y}_{\Lambda_{M-1,\|}}^\dagger\mathbf{y}_{N,\|}\|_2 = \|\mathrm{P}_{\mathcal{B}}\mathbf{y}_{N,\|}\|_2,\end{aligned} \tag{56}$$

where $\mathcal{B}=\mathcal{S}_L\cap\mathcal{R}(\mathbf{Y}_{\Lambda_{M-1,\|}})^\perp$ is a subspace of dimension $d_L-p(M-1)>0$. Since $\mathbf{y}_{N,\|}=\mathbf{x}_N+\mathbf{e}_{N,\|}$ and both the distributions of $\mathbf{x}_N$ and $\mathbf{e}_{N,\|}$ are rotationally invariant in $\mathcal{S}_L$ (Tschannen & Bölcskei, 2018, p. 4092), the normalized $\mathbf{y}_{N,\|}/\|\mathbf{y}_{N,\|}\|_2$ is uniformly distributed over $\mathcal{S}_L\cap\mathbb{S}^{n-1}$. Let $\mathbf{V}_{\mathcal{B}}=[\mathbf{v}_1\ \mathbf{v}_2...\mathbf{v}_{d_L-p(M-1)}]\in\mathbb{R}^{n\times(d_L-p(M-1))}$ be a matrix whose columns form an orthonormal basis for $\mathcal{B}$; augment $\mathbf{V}_{\mathcal{B}}$ by adding extra $p(M-1)$ columns so that the d_L columns of $\mathbf{V}_{\mathcal{B}}=[\mathbf{v}_1\ \mathbf{v}_2...\mathbf{v}_{d_L-p(M-1)}...\mathbf{v}_{d_L}]\in\mathbb{R}^{n\times d_L}$ form an orthonormal

basis for $\mathcal{S}_L$. Clearly, we have $\mathbf{y}_{N,\|}/\|\mathbf{y}_{N,\|}\|_2 = \mathbf{V}\mathbf{a}$, where $\mathbf{a} = [a_1\ a_2 ... a_{d_L}]^T$ obeys uniform distribution over the unit sphere of $\mathbb{R}^{d_L}$ and, in particular, $\mathrm{P}_{\mathcal{B}}(\mathbf{y}_{N,\|}/\|\mathbf{y}_{N,\|}\|_2) = \mathbf{V}_{\mathcal{B}}\widetilde{\mathbf{a}}$, where $\widetilde{\mathbf{a}} = [a_1\ a_2 ... a_{d_L-p(M-1)}]^T$ obeys the distribution specified in (Knokhlov, 2006, eq. (7)). Noticing $\|\mathrm{P}_{\mathcal{B}}(\mathbf{y}_{N,\|}/\|\mathbf{y}_{N,\|}\|_2)\|_2 = \|\mathbf{V}_{\mathcal{B}}\widetilde{\mathbf{a}}\|_2 = \|\widetilde{\mathbf{a}}\|_2$, the following set of inequalities hold

$$
\begin{aligned}
&\Pr\Big\{\Big\|\mathrm{P}_{\mathcal{B}}\Big(\frac{\mathbf{y}_{N,\|}}{\|\mathbf{y}_{N,\|}\|_2}\Big)\Big\|_2 \le \lambda\Big\} = \Pr\{\|\widetilde{\mathbf{a}}\|_2 \le \lambda\} \\
&\overset{(a)}{\le} \Big(\frac{\Gamma(d_L/2)}{\pi^{(d_L-p(M-1))/2}\Gamma(p(M-1)/2)}\Big)\int_{\|\widetilde{\mathbf{a}}\|_2\le\lambda}(1-\|\widetilde{\mathbf{a}}\|_2^2)^{(p(M-1)-2)/2}d\widetilde{\mathbf{a}} \\
&\overset{(b)}{\le} \Big(\frac{d_L}{2\pi}\Big)^{(dL-p(M-1))/2}\int_{\|\widetilde{\mathbf{a}}\|_2\le\lambda}(1-\|\widetilde{\mathbf{a}}\|_2^2)^{(p(M-1)-2)/2}d\widetilde{\mathbf{a}} \\
&\le \Big(\frac{d_L}{2\pi}\Big)^{(dL-p(M-1))/2}\int_{\|\widetilde{\mathbf{a}}\|_2\le\lambda} 1\ d\widetilde{\mathbf{a}} \\
&= \Big(\frac{d_L}{2\pi}\Big)^{(dL-p(M-1))/2}\ v(d_L-p(M-1))\lambda^{d_L-p(M-1)} \\
&\overset{(c)}{\le} \Big(\frac{d_L}{2\pi}\Big)^{(dL-p(M-1))/2}\ 6\lambda^{d_L-p(M-1)},
\end{aligned}
\tag{57}
$$

where (a) holds by (Knokhlov, 2006, eq. (7)), (b) follows from (Foucart & Rauhut, 2013, eq. (8.1)), in which $v(r)$ is the volume of unit-ball in $\mathbb{R}^r$, and (c) is true since $v(r) \le 6$ for all $r \in \mathbb{N}$ (Smith & Vamanamurthy, 1989). Using (56) and (57) with $\lambda = \sigma\sqrt{2/d_L}$, we can obtain

$$
\|\mathbf{r}_{m-1,\|}^{(N)}\|_2 > \sqrt{\frac{2}{d_L}}\sigma\|\mathbf{y}_{N,\|}\|_2,\ 1 \le m \le M, \tag{58}
$$

holds with a probability at least $1 - 6(\sigma/\sqrt{\pi})^{d_L-p(M-1)}$. Since $\mathbf{y}_{N,\|} = \mathbf{x}_N + \mathbf{e}_{N,\|}$ and $\|\mathbf{x}_N\|_2 = 1$ (see Assumption 1), triangle inequality gives $\|\mathbf{y}_{N,\|}\|_2 \ge 1 - \|\mathbf{e}_{N,\|}\|_2$, which together with (58) implies

$$
\|\mathbf{r}_{m-1,\|}^{(N)}\|_2 > \sqrt{\frac{2}{d_L}}\sigma(1-\|\mathbf{e}_{N,\|}\|_2),\ 1 \le m \le M, \tag{59}
$$

holds with a probability as high as (55). With the aid of (53), (54), and (59), we go on to find a sufficient condition ensuring (51) by finding an upper bound for the LHS of (51). Once (53), (54), and (59) hold, we can obtain the following set of inequalities

$$
\begin{aligned}
&\max_{l\ne L}\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l}} + \frac{\sqrt{3d_L}\|\mathbf{r}_{m,\perp}^{(N)}\|_2}{\sqrt{8(n-d_L)\log N}\|\mathbf{r}_{m,\|}^{(N)}\|_2} + \frac{2\sqrt{3d_L}(1+\|\mathbf{e}_N\|_2)}{\sqrt{8n\log N}\|\mathbf{r}_{m-1,\|}^{(N)}\|_2}\max_{1\le i\le N}\|\mathbf{e}_i\|_2 \\
&\le \max_{l\ne L}\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l}} + \frac{3\sqrt{3}d_L(3+3\sigma)}{(8-12\sigma)\sqrt{n\log N}} \\
&\overset{(a)}{\le} \max_{l\ne L}\ aff(\mathcal{S}_l,\mathcal{S}_L) + \frac{3\sqrt{3}d_L(3+3\sigma)}{(8-12\sigma)\sqrt{n\log N}} \overset{(b)}{\le} \frac{\tau}{4\log N},
\end{aligned}
\tag{60}
$$

where (a) follows from the definition (9) and (b) is true thanks to Assumption 3 As a result, when (53), (54) and (59) hold, the inequality in (60) for all $1 \le m \le M$ is guaranteed. By employing the union bound technique it can be concluded that the inequality (51) for all $1 \le m \le M$ holds with a probability at least $1 - 6(\sigma/\sqrt{\pi})^{d_L-p(M-1)} - Ne^{-n/8}$. Finally, by using (36), (39), (44), (45), (49) and (51), the proof is thus completed again by using the union bound. □

5.2 Proof of Corollary 1

The lower bound (13) is simply obtained as the probability of the event that the numbers k_m's of recovered true neighbors throughout all M iterations yield the maximal lower bound (11). That is to say, (13) is the

maximum of (11) over all feasible $0 \leq k_m \leq p$, $1 \leq m \leq M$; more precisely, the minimum objective of the following optimization problem

$$\begin{aligned} &\min_{(k_1,...,k_m)} J_0 = \sum_{m=1}^{M} J(k_m) \\ &\text{s.t.} \ \sum_{m=1}^{M} k_m - k_t = 0 \\ &\quad k_m \in \{0,1,2,...,p\}, \ \forall 1 \leq m \leq M, \end{aligned} \tag{61}$$

where

$$\begin{aligned} J(k_m) \triangleq \Bigg(&\left(\frac{2e(N-|\mathcal{Y}_L|)}{(p-k_m+1)N^{8\log N/d_L}}\right)^{p-k_m+1} \\ &+ \left(\sqrt{\frac{2}{\pi}}\tau\right)^{|\mathcal{Y}_L|-d_L-k_m} \left(\frac{e(|\mathcal{Y}_L|-1)}{k_m-1}\right)^{k_m-1} + \frac{4+2c}{N^2}\Bigg)\mathbf{1}(k_m>0) \end{aligned} \tag{62}$$

Below, we will show that the two-level sequence

$$(k_1,...,k_m) = \Big(\underbrace{q_t+1, q_t+1, ..., q_t+1}_{r_t-\text{fold}}, \underbrace{q_t, q_t, ..., q_t}_{(M-r_t)-\text{fold}}\Big) \tag{63}$$

solves (61), consequently leading to (13). We recall the following definition and lemma, which are needed in our proof.

Definition 1. *(Pečarić et al., 1992) Let* $\mathbf{x} = [x_1 \ x_2 ... x_M]^T \in \mathbb{R}^M$ *and* $\mathbf{y} = [y_1 \ y_2 ... y_M]^T \in \mathbb{R}^M$ *be two real vectors whose entries are ordered in the way that* $x_{[1]} \geq x_{[2]} \geq ... \geq x_{[M]}$ *and* $y_{[1]} \geq y_{[2]} \geq ... \geq y_{[M]}$*, respectively. Then* $\mathbf{x}$ *is said to be majorized by* $\mathbf{y}$ *if*

$$\sum_{m=1}^{s} x_{[m]} \leq \sum_{m=1}^{s} y_{[m]}, \forall 1 \leq s \leq M \tag{64}$$

and

$$\sum_{m=1}^{s} x_m = \sum_{m=1}^{s} y_m. \tag{65}$$

We say $f : \mathcal{A} \subset \mathbb{R}^M$ *is Schur-convex if* $f(\mathbf{x}) \leq f(\mathbf{y})$ *whenever* $\mathbf{x}$ *is majorized by* $\mathbf{y}$*,* $\mathbf{x}, \mathbf{y} \in \mathcal{A}$*.* □

Lemma 4. *(Pečarić et al., 1992, Theorem 12.25) Let* $f : \mathcal{A} \subset \mathbb{R}^M$ *be a permutation invariant function, that is,* $f(\mathbf{x}) = f(\mathbf{P}\mathbf{x})$ *for all* $\mathbf{x} \in \mathcal{A}$ *and permutation matrices* $\mathbf{P} \in \mathbb{R}^{M\times M}$*, whose first partial derivatives exist in* $\mathcal{A}$*. Then* f *is Schur-convex in* $\mathcal{A}$ *if and only if*

$$(x_i - x_j)\left(\frac{\partial f}{\partial x_i} - \frac{\partial f}{\partial x_j}\right) \geq 0, \forall \mathbf{x} = [x_1 \ x_2 ... x_M]^T \in \mathcal{A} \tag{66}$$

holds for all $1 \leq i \neq j \leq M$*.* □

Using vector-matrix notation, all we have to do is to show

$$\mathbf{k}^* \triangleq \Big[\underbrace{q_t+1, q_t+1, ..., q_t+1}_{r_t-\text{fold}}, \underbrace{q_t, q_t, ..., q_t}_{(M-r_t)-\text{fold}}\Big]^T \tag{67}$$

solves the following optimization problem:

$$\begin{aligned} &\min_{\mathbf{k}=[k_1,...,k_m]^T} J_0 = \sum_{m=1}^{M} J(\mathbf{k}^T\mathbf{v}_m) \\ &\text{s.t.} \ \mathbf{k}^T\mathbf{1} = k_t \\ &\quad \mathbf{k}^T\mathbf{v}_m \in \{0,1,2,...,p\}, \ \forall 1 \leq m \leq M, \end{aligned} \tag{68}$$

in which $\mathbf{v}_m \triangleq [\ \underbrace{0...0}_{m-1 \text{ fold}}\ \ 1\ 0...0\ ]^T$ is the mth standard unit vector. It suffices to prove

$$J_0(\mathbf{k}^*) \leq J_0(\mathbf{k}) \tag{69}$$

for any $\mathbf{k} \in \mathcal{D}_J$, the feasible set of (68), based on Lemma 4. For this we first note that the objective J_0 is not differentiable, since the function J in (62) involves the indicator function. To rid of this difficulty, we consider the differentiable surrogate $\widetilde{J} : \mathbb{R} \to \mathbb{R}$ for J, constructed according to

$$\widetilde{J}(x) = J(x),\ \forall x \geq 0.9 \text{ and } x = 0, \tag{70}$$

and

$$\widetilde{J}'(0) = J'(1). \tag{71}$$

Thanks to (70), the function $\widetilde{J}$ thus obtained satisfies $\widetilde{J}(k_m) = J(k_m)$ for all $k_m \in \{0, 1, ..., p\}$ and $\widetilde{J}'(k_m) = J'(k_m)$ for all $k_m \in \{1, 2, ..., p\}$. The corresponding differentiable surrogate for J_0 is accordingly given by

$$\widetilde{J}_0(\mathbf{k}) = \sum_{m=1}^{M} \widetilde{J}(\mathbf{k}^T \mathbf{v}_m). \tag{72}$$

Clearly, $\widetilde{J}_0(\mathbf{k}) = J_0(\mathbf{k})$ for all $\mathbf{k} \in \mathcal{D}_J$, due to condition (70); all the better, $\widetilde{J}_0$ is Schur-convex in $\mathcal{D}_J$ (a proof is given in Appendix C). Hence, for $\mathbf{k}, \mathbf{q} \in \mathcal{D}_J$ such that $\mathbf{k}$ is majorized by $\mathbf{q}$, we have $J_0(\mathbf{k}) = \widetilde{J}_0(\mathbf{k}) \leq \widetilde{J}_0(\mathbf{q}) = J_0(\mathbf{q})$. The inequality (69) is guaranteed once $\mathbf{k}^*$ is majorized by any feasible $\mathbf{k}$, which is indeed true as shown below. For a feasible $\mathbf{k} = [k_1\ k_2 ... k_M]^T$ such that $k_1 \geq k_2 \geq ... \geq k_M$ without loss of generality. By Definition 1, it suffices to show

$$\sum_{m=1}^{s} k_m^* \leq \sum_{m=1}^{s} k_m,\ \forall 1 \leq s \leq M. \tag{73}$$

Assume otherwise that there exists $1 \leq s \leq M$ such that $\sum_{m=1}^{q} k_m^* \leq \sum_{m=1}^{q} k_m$, $1 \leq q \leq s-1$, whereas $\sum_{m=1}^{s} k_m^* > \sum_{m=1}^{s} k_m$. Then we have $k_s^* > k_s$, which together with $k_s^* \in \{q_t, q_t + 1\}$ implies $q_t \geq k_m$, for $s + 1 \leq m \leq M$, ending in the following contradiction:

$$\begin{aligned} k_t = \sum_{m=1}^{s} k_m + \sum_{m=s+1}^{M} k_m \leq \sum_{m=1}^{s} k_m + \sum_{m=s+1}^{M} q_t \leq \sum_{m=1}^{s} k_m + \sum_{m=s+1}^{M} k_m^* \\ < \sum_{m=1}^{s} k_m^* + \sum_{m=s+1}^{M} k_m^* = k_t. \end{aligned} \tag{74}$$

□

5.3 Proof of Theorem 2

Below we first derive an equivalent condition for the proposed stopping rule that is more amenable to analysis. Recall the residual $\mathbf{r}_m^{(N)}$ obtained in the mth iteration is the orthogonal projection of the previous residual $\mathbf{r}_{m-1}^{(N)}$ onto $\mathcal{R}(\mathbf{Y}_{\Lambda_m})^\perp$. Since $\|\mathbf{r}_{m-1}^{(N)}\|_2 = \|\mathbf{r}_m^{(N)}\|_2 + \|\mathbf{r}_{m-1}^{(N)} - \mathbf{r}_m^{(N)}\|_2$, obtained from the Pythagorean theorem, the proposed stopping rule can be rewritten as

$$1 - \frac{\|\mathbf{r}_m^{(N)}\|_2}{\|\mathbf{r}_{m-1}^{(N)}\|_2} = 1 - \sqrt{1 - \|\widetilde{\mathbf{r}}_m^{(N)}\|_2^2} \leq \sqrt{\frac{p}{n}}, \tag{75}$$

where

$$\widetilde{\mathbf{r}}_m^{(N)} \triangleq \frac{\mathbf{r}_{m-1}^{(N)} - \mathbf{r}_m^{(N)}}{\|\mathbf{r}_{m-1}^{(N)}\|_2} \in \mathcal{R}(\mathbf{Y}_{\Lambda_m}) \tag{76}$$

is the normalized difference of the residual vectors. Rearranging the inequality in (75) yields the following equivalent stopping condition

$$\|\widetilde{\mathbf{r}}_m^{(N)}\|_2 \leq \sqrt{2\sqrt{p/n} - p/n}. \tag{77}$$

Below we show that, with a high chance, (77) does not hold when $m \leq \lceil d_L/p \rceil$ and is achieved (hence, GOMP stops) with $m = \lceil d_L/p \rceil + 1$. Formally, we prove the following inequalities

$$\|\widetilde{\mathbf{r}}_m^{(N)}\|_2 > \sqrt{2\sqrt{p/n} - p/n}, \ \forall 1 \leq m \leq \lceil d_L/p \rceil \tag{78}$$

and

$$\|\widetilde{\mathbf{r}}_{\lceil d_L/p \rceil+1}^{(N)}\|_2 \leq \sqrt{2\sqrt{p/n} - p/n} \tag{79}$$

hold at once with a probability as high as claimed by Theorem 2.

To proceed, let $\mathbf{y}_j$ be a selected data vector in the mth iteration, $1 \leq m \leq \lceil d_L/p \rceil$. As long as (45), (49), (53), (54) and (59) hold, we have

$$\begin{aligned}
|\langle \mathbf{y}_j, \mathbf{r}_{m-1}^{(N)} \rangle| &\overset{(a)}{\geq} |\langle \mathbf{y}_{k_m}, \mathbf{r}_{m-1}^{(N)} \rangle| \overset{(b)}{\geq} |\langle \mathbf{x}_{\bar{t}_{k_m}}, \mathbf{r}_{m-1,\|}^{(N)} \rangle| - \max_{1 \leq j \leq |\mathcal{Y}_L| - 1 - |\Lambda_{m-1}^t|} |\langle \mathbf{e}_{t_j}, \mathbf{r}_{m-1}^{(N)} \rangle| \\
&\overset{(c)}{\geq} \Big[\frac{\|\mathbf{r}_{m-1,\|}^{(N)}\|_2}{\sqrt{d_L}} - \sqrt{\frac{6 \log N}{n}} \|\mathbf{r}_{m-1,\|}^{(N)}\|_2 \max_{1 \leq i \leq N} \|\mathbf{e}_i\|_2 \Big] \\
&= \|\mathbf{r}_{m-1}^{(N)}\|_2 \Big[\frac{\|\mathbf{r}_{m-1,\|}^{(N)}\|_2}{\sqrt{d_L(\|\mathbf{r}_{m-1,\perp}^{(N)}\|_2^2 + \|\mathbf{r}_{m-1,\|}^{(N)}\|_2^2)}} - \sqrt{\frac{6 \log N}{n}} \max_{1 \leq i \leq N} \|\mathbf{e}_i\|_2 \Big] \\
&\overset{(d)}{\geq} \|\mathbf{r}_{m-1}^{(N)}\|_2 \underbrace{\Big[\frac{\sqrt{2}(1 - 3\sigma/2)}{\sqrt{9d_L^2/4 + 4(1 - 3\sigma/2)^2 d_L}} - \frac{3\sqrt{6 \log N}\sigma}{2\sqrt{n}} \Big]}_{\triangleq \eta},
\end{aligned} \tag{80}$$

in which (a) is true since $\mathbf{y}_j$ is selected as a neighbor, (b) follows form (31), (c) holds due to (45) and (49), and (d) is obtained by combining (53), (54), and (59). By definition (76) we have $\widetilde{\mathbf{r}}_m^{(N)} = \mathrm{P}_{\mathcal{R}(\mathbf{Y}_{\Lambda_m})}(\mathbf{r}_{m-1}^{(N)}/\|\mathbf{r}_{m-1}^{(N)}\|_2)$, hence

$$\|\widetilde{\mathbf{r}}_m^{(N)}\|_2 \geq \|\mathrm{P}_{span\{\mathbf{y}_j\}}(\frac{\mathbf{r}_{m-1}^{(N)}}{\|\mathbf{r}_{m-1}^{(N)})\|_2})\|_2 = \frac{|\langle \mathbf{y}_j, \mathbf{r}_{m-1}^{(N)} \rangle|}{\|\mathbf{y}_j\|_2 \|\mathbf{r}_{m-1}^{(N)}\|_2}, \tag{81}$$

which together with (80) implies

$$\|\widetilde{\mathbf{r}}_m^{(N)}\|_2 \geq \frac{|\langle \mathbf{y}_j, \mathbf{r}_{m-1}^{(N)} \rangle|}{\|\mathbf{y}_j\|_2 \|\mathbf{r}_{m-1}^{(N)}\|_2} \geq \frac{\eta \|\mathbf{r}_{m-1}^{(N)}\|_2}{\|\mathbf{y}_j\|_2 \|\mathbf{r}_{m-1}^{(N)}\|_2} = \frac{\eta}{\|\mathbf{y}_j\|_2}. \tag{82}$$

With the aid of (82), inequality (78) is guaranteed once $\eta/\|\mathbf{y}_j\|_2 > \sqrt{2\sqrt{p/n} - p/n}$, which is typically true since the ambient dimension n is drastically large.

As for (79), again by definition (76) we write

$$\widetilde{\mathbf{r}}_{\lceil d_L/p \rceil+1}^{(N)} = \mathrm{P}_{\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil+1}})}(\mathbf{r}_{\lceil d_L/p \rceil}^{(N)}/\|\mathbf{r}_{\lceil d_L/p \rceil}^{(N)}\|_2). \tag{83}$$

Let $\{\mathbf{b}_1, ..., \mathbf{b}_{(\lceil d_L/p \rceil+1)p}\}$ be an orthonormal basis of $\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil+1}})$, arranged in a way that $\{\mathbf{b}_{p+1}, ..., \mathbf{b}_{(\lceil d_L/p \rceil+1)p}\}$ is an orthonormal basis of $\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}})$. Since $\mathbf{r}_{\lceil d_L/p \rceil}^{(N)}$ is orthogonal to $\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}})$,

$\langle \mathbf{r}^{(N)}_{\lceil d_L/p \rceil}, \mathbf{b}_j \rangle = 0$ for all $p+1 \leq j \leq (\lceil d_L/p \rceil + 1)p$. Thus, it follows

$$\begin{aligned}
&\Pr\{\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil+1}\|_2 \leq \sqrt{2\sqrt{p/n} - p/n}\} \\
&= \Pr\Big\{\|\mathrm{P}_{\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil+1}})}(\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2)\|_2 \leq \sqrt{2\sqrt{p/n} - p/n}\ \Big\} \\
&\geq \Pr\Big\{\bigcap_{k=1}^{p}\Big\{\|\mathrm{P}_{span\{\mathbf{b}_k\}}(\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2)\|_2 \leq \sqrt{2\sqrt{p/n} - p/n}/\sqrt{p}\Big\}\Big\} \\
&= 1 - \Pr\Big\{\bigcup_{k=1}^{p}\Big\{\|\mathrm{P}_{span\{\mathbf{b}_k\}}(\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2)\|_2 > \sqrt{2\sqrt{p/n} - p/n}/\sqrt{p}\Big\}\Big\} \\
&\geq 1 - \sum_{k=1}^{p}\Pr\Big\{\|\mathrm{P}_{span\{\mathbf{b}_k\}}(\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2)\|_2 > \sqrt{2\sqrt{p/n} - p/n}/\sqrt{p}\Big\} \\
&= 1 - \sum_{k=1}^{p}\Pr\Big\{|\langle \mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2, \mathbf{b}_k\rangle| > \sqrt{2\sqrt{p/n} - p/n}/\sqrt{p}\Big\}.
\end{aligned} \tag{84}$$

According to our proof of Theorem 1, the neighbors selected in all $\lceil d_L/p \rceil$ iterations (i.e., all columns of $\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}}$) are correct when (36), (39), (44), (45), (49), (53), (54) and (59) hold for all $1 \leq m \leq \lceil d_L/p \rceil$. If so, there then exists $\widetilde{\mathbf{c}} \in \mathbb{R}^{p\lceil d_L/p \rceil}$ such that the "noiseless" signal vector $\mathbf{x}_N$ and neighbor matrix $\mathbf{X}_{\Lambda_{\lceil d_L/p \rceil}}$ satisfy

$$\mathbf{x}_N = \mathbf{X}_{\Lambda_{\lceil d_L/p \rceil}}\widetilde{\mathbf{c}}. \tag{85}$$

Noting that $\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}} = \mathbf{X}_{\Lambda_{\lceil d_L/p \rceil}} + \mathbf{E}_{\Lambda_{\lceil d_L/p \rceil}}$, where $\mathbf{E}_{\Lambda_{\lceil d_L/p \rceil}}$ is the noise matrix, the residual vector at the $\lceil d_L/p \rceil$th iteration can be written as

$$\begin{aligned}
\mathbf{r}^{(N)}_{\Lambda_{\lceil d_L/p \rceil}} &= (\mathbf{I} - \mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}}\mathbf{Y}^{\dagger}_{\Lambda_{\lceil d_L/p \rceil}})(\mathbf{x}_N + \mathbf{e}_N) \\
&= (\mathbf{I} - \mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}}\mathbf{Y}^{\dagger}_{\Lambda_{\lceil d_L/p \rceil}})(\mathbf{x}_N + \mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}}\widetilde{\mathbf{c}} + \mathbf{e}_N) \\
&= (\mathbf{I} - \mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}}\mathbf{Y}^{\dagger}_{\Lambda_{\lceil d_L/p \rceil}})(\mathbf{E}_{\Lambda_{\lceil d_L/p \rceil}}\widetilde{\mathbf{c}} + \mathbf{e}_N),
\end{aligned} \tag{86}$$

which is the projection of the Gaussian random vector $\mathbf{E}_{\Lambda_{\lceil d_L/p \rceil}}\widetilde{\mathbf{c}} + \mathbf{e}_N$ onto $\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}})^{\perp}$. Being a linear combination of rotationally invariant vectors, $\mathbf{E}_{\Lambda_{\lceil d_L/p \rceil}}\widetilde{\mathbf{c}} + \mathbf{e}_N$ remains so; this implies the projection $\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}$ is also rotationally invariant on $\mathcal{R}(\mathbf{Y}_{\Lambda_{\lceil d_L/p \rceil}})^{\perp}$. Then using (24) with $\mathbf{a} = \mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2$ and $\mathbf{b} = \mathbf{b}_k$, $1 \leq k \leq p$, we can obtain

$$\Pr\Big\{|\langle \mathbf{r}^{(N)}_{\lceil d_L/p \rceil}/\|\mathbf{r}^{(N)}_{\lceil d_L/p \rceil}\|_2, \mathbf{b}_k\rangle| > \sqrt{2\sqrt{p/n} - p/n}/\sqrt{p}\Big\} \approx 2e^{-\sqrt{n/p}}, \tag{87}$$

which together with (84) implies (79) holds with a probability at least $1 - 2pe^{-\sqrt{n/p}}$. In summary, once the inequalities (36), (39), (44), (45), (49), (53), (54), and (59) hold for all $1 \leq m \leq \lceil d_L/p \rceil$, then (78) and (79) hold with a probability at least $1 - 2pe^{-\sqrt{n/p}}$. The proof is then completed by employing the union bound.

□

6 Conclusion

Fast acquisition of many true neighbors underlies the success of SSC in real-world applications. Under the framework of greedy selection, which is pretty suited for low-complexity implementation, we propose a GOMP sparse regression scheme, together with a novel subspace-dimensionality-aware stopping rule, to boost neighbor recovery. Thanks to multiple neighbor identification per iteration, the proposed GOMP involves fewer iterations, thereby enjoying even lower algorithmic complexity than conventional OMP; in

addition, the residual vector better stands up to noise corruption, consequently bringing about higher neighbor identification accuracy. Our proposed stopping criterion is appealing in that it depends entirely on a knowledge of the ambient space dimension, in marked contrast with the existing solution (Tschannen & Bölcskei, 2018) that requires extra off-line estimation of either subspace dimension or noise power. Besides algorithm development, in-depth neighbor recovery rate analyses were conducted to justify the merits of the proposed GOMP; the obtained analytic results are further validated by computations using both synthetic and real datasets. Overall, our study presents a new greedy-based SSC scheme, with provable performance guarantees, that can pave the way for practical applications. Future work will analyze AoD of both GOMP and OMP under the considered semi-random model, particularly to show GOMP enjoys a smaller average AoD; an analytic study of AoD would further offer certain guidelines on determining the optimum number of neighbors to recover in each iteration, for the purpose of improving the current scheme, which identifies a constant number ($p \geq 1$) of neighbors throughout all iterations, in a more dynamic environment.

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A Appendix

Assume $\mathbf{x}_i \in \mathcal{S}_l$, $l \neq L$. Then we have

$$\begin{aligned}\Pr\left\{\|\mathbf{U}_L^T\mathbf{x}_i\|_2 > 4\log N\frac{\|\mathbf{U}_L^T\mathbf{U}_l\|_F}{\sqrt{d_L d_l}}\right\} &\overset{(a)}{=} \Pr\left\{\|\mathbf{U}_L^T\mathbf{x}_i\|_2 > \frac{4\log N}{\sqrt{d_L}}\sqrt{E_u[\|\mathbf{U}_L^T\mathbf{U}_l\mathbf{a}\|_2^2]}\right\} \\ &\overset{(b)}{\leq} \frac{2}{N^{(8\log N)/d_L}},\end{aligned} \tag{88}$$

where (a) follows from (Soltanolkotabi & Candès, 2012, p. 2231) , and (b) holds by the same procedures for the proof of (Soltanolkotabi & Candès, 2012, Lemma 7.5) . Since $\mathbf{r}_{m,\|}^{(N)} \in \mathcal{S}_L$, we have $\mathbf{r}_{m,\|}^{(N)}/\|\mathbf{r}_{m,\|}^{(N)}\|_2 = \mathbf{U}_L\mathbf{z}$ with $\|\mathbf{z}\|_2 = 1$. The Cauchy-Schwartz inequality implies

$$\left|\left\langle \mathbf{r}_{m,\|}^{(N)}/\|\mathbf{r}_{m,\|}^{(N)}\|_2, \mathbf{x}_i\right\rangle\right| = |\mathbf{z}^T\mathbf{U}_L^T\mathbf{x}_i| \leq \|\mathbf{z}\|_2\|\mathbf{U}_L^T\mathbf{x}_i\|_2 = \|\mathbf{U}_L^T\mathbf{x}_i\|_2, \tag{89}$$

and therefore

$$\left\{\left|\left\langle \mathbf{x}_i, \mathbf{r}_{m,\|}^{(N)}/\|\mathbf{r}_{m,\|}^{(N)}\|_2\right\rangle\right| > 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_L d_l}}\right\} \subset \left\{\|\mathbf{U}_L^T\mathbf{x}_i\|_2 > 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_L d_l}}\right\}. \tag{90}$$

With (88) and (90), it follows immediately

$$\Pr\left\{\left|\left\langle \mathbf{x}_i, \mathbf{r}_{m,\|}^{(N)}/\|\mathbf{r}_{m,\|}^{(N)}\|_2\right\rangle\right| > 4\log(N)\frac{\|\mathbf{U}_l^T\mathbf{U}_L\|_F}{\sqrt{d_l d_L}}\right\} < \frac{2}{N^{(8\log N)/d_L}}. \tag{91}$$

□

Appendix B

Let

$$\mathcal{D}(\mathbf{v}) \triangleq \Big\{(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) : \mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2 = \mathbf{v}\Big\}, \tag{92}$$

$f : \mathbb{R}^n \times ... \times \mathbb{R}^n \to \mathbb{R}$ be the probability density function of $(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N)$, and $\mathbb{S}^{n-1}$ be the unit-sphere of $\mathbb{R}^n$. Then for $\mathbf{z} \in \mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}$, we have

$$\begin{aligned}\Pr\Big\{|\langle \mathbf{z}, \mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2\rangle| > \epsilon\Big\} &= \int\limits_{(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r}:|\langle \mathbf{r},\mathbf{z}\rangle|>\epsilon\}} \Big[\int\limits_{\mathcal{D}(\mathbf{v})} f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) d\mathbf{y}_1 d\mathbf{y}_2 ... d\mathbf{y}_N\Big] d\mathbf{v} \\ &\leq \int\limits_{(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r}:|\langle \mathbf{r},\mathbf{z}\rangle|\geq\epsilon\}} \Big[\int\limits_{\mathcal{K}\cap\mathcal{D}(\mathbf{v})} f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) d\mathbf{y}_1 d\mathbf{y}_2 ... d\mathbf{y}_N \\ &\qquad + \int\limits_{\mathcal{K}^c\cap\mathcal{D}(\mathbf{v})} f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) d\mathbf{y}_1 d\mathbf{y}_2 ... d\mathbf{y}_N\Big] d\mathbf{v},\end{aligned} \tag{93}$$

where the last inequality holds by defining

$$\mathcal{K} \triangleq \{\mathbf{x} + \mathbf{e} : \mathbf{x} \in \mathcal{S}_{k_1} \cap \mathbb{S}^{n-1}, \|\mathbf{e}\|_2 \leq 3\sigma/2\} \times ... \times \{\mathbf{x} + \mathbf{e} : \mathbf{x} \in \mathcal{S}_{k_N} \cap \mathbb{S}^{n-1}, \|\mathbf{e}\|_2 \leq 3\sigma/2\}, \tag{94}$$

in which we assume that $\mathbf{y}_i \in \mathcal{Y}_{k_i}$, hence $\mathbf{x}_i \in \mathcal{S}_{k_i}$, for $1 \leq i \leq N$. Since the maximum function is continuous and composition of continuous functions is also continuous, $\mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2$ is a continuous function of $(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N)$. Consequently, since $(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r} : |\langle \mathbf{r}, \mathbf{z}\rangle| \geq \epsilon\}$ is closed, the inverse image of $\mathcal{D}((\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r} : |\langle \mathbf{r}, \mathbf{z}\rangle| \geq \epsilon\})$ is also closed (Marsden & Hoffman, 1993). Notice that $\mathcal{K} \cap \mathcal{D}((\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r} : |\langle \mathbf{r}, \mathbf{z}\rangle| \geq \epsilon\})$ is closed and bounded in the Euclidean space, and is therefore compact by the Heine-Borel theorem (Marsden & Hoffman, 1993). Since $f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N)$ is continuous on the compact set $\mathcal{K} \cap \mathcal{D}((\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r} : |\langle \mathbf{r}, \mathbf{z}\rangle| \geq \epsilon\})$, by the extreme value theorem (Marsden & Hoffman, 1993) there exists a constant δ such that

$$f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) \leq \delta, \ \forall(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) \in \mathcal{K} \cap \mathcal{D}((\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r} : |\langle \mathbf{r}, \mathbf{z}\rangle| \geq \epsilon\}), \tag{95}$$

which together with (93) implies

$$\begin{aligned}\Pr\Big\{\Big|\Big\langle \mathbf{z}, \frac{\mathbf{r}_{m,\perp}^{(N)}}{\|\mathbf{r}_{m,\perp}^{(N)}\|_2}\Big\rangle\Big| > \epsilon\Big\} &= \int\limits_{(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r}:|\langle \mathbf{r},\mathbf{z}\rangle|\geq\epsilon\}} \Big[\delta|\mathcal{K}| + \int\limits_{\mathcal{K}^c\cap\mathcal{D}(\mathbf{v})} f(\mathbf{y}_1, \mathbf{y}_2, ..., \mathbf{y}_N) d\mathbf{y}_1 d\mathbf{y}_2 ... d\mathbf{y}_N\Big] d\mathbf{v} \\ &\overset{(a)}{\leq} \int\limits_{(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r}:|\langle \mathbf{r},\mathbf{z}\rangle|\geq\epsilon\}} \Big[\delta|\mathcal{K}| + Ne^{-n/8}\Big] d\mathbf{v} \\ &\overset{(b)}{=} \int\limits_{(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}) \cap \{\mathbf{r}:|\langle \mathbf{r},\mathbf{z}\rangle|\geq\epsilon\}} \Big[c/A(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1})\Big] d\mathbf{v} \overset{(c)}{\leq} 2ce^{-(n-d_L)\epsilon^2/2},\end{aligned} \tag{96}$$

where (a) holds by (Tschannen & Bölcskei, 2018, Lemma 10) with $|\mathcal{K}|$ here denoting the Lebesgue measure of $\mathcal{K}$, (b) is true as we define $A(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1})$ to be the surface area of $\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1}$ and the constant c is defined as $c \triangleq (\delta|\mathcal{K}| + Ne^{-n/8})A(\mathcal{S}_L^\perp \cap \mathbb{S}^{n-1})$, and (c) follows from (24). With (96), we reach

$$\begin{aligned}\Pr\Big\{\Big|\Big\langle \mathbf{x}_i, \mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2\Big\rangle\Big| > \epsilon\Big\} &= \int\limits_{\mathbb{R}^n} \Pr\Big\{\Big|\Big\langle \mathbf{x}_i, \mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2\Big\rangle\Big| > \epsilon \Big| \mathbf{x}_i = \mathbf{v}\Big\} f_{\mathbf{x}_i}(\mathbf{v}) d\mathbf{v} \\ &\overset{(a)}{\leq} \int\limits_{\mathbb{R}^n} 2en^{-(n-d_L)\epsilon^2/2} f_{\mathbf{x}_i}(\mathbf{v}) d\mathbf{v} = 2en^{-(n-d_L)\epsilon^2/2},\end{aligned} \tag{97}$$

where (a) holds by (96), in which $f_{\mathbf{x}_i}$ is the probability density function of $\mathbf{x}_i$. Let $\epsilon = \sqrt{6\log N/(n-d_L)}$. Then (97) gives

$$\Pr\Big\{\Big|\Big\langle \mathbf{x}_i, \mathbf{r}_{m,\perp}^{(N)}/\|\mathbf{r}_{m,\perp}^{(N)}\|_2\Big\rangle\Big| > \sqrt{\frac{6\log N}{n-d_L}}\Big\} \leq 2cN^{-3}. \tag{98}$$

□

Appendix C

In this appendix, we prove the Schur-convexity of $\widetilde{J}_0$. Notably, straightforward manipulations show

$$(k_m - k_q)\Big(\frac{\partial J_0}{\partial k_m} - \frac{\partial J_0}{\partial k_q}\Big) = (k_m - k_q)(J'(k_m) - J'(k_q)) \geq 0,\ \forall k_m, k_q \in \{1, 2, ..., p\}. \tag{99}$$

It then follows

$$\begin{aligned}(k_m - k_q)\Big(\frac{\partial \widetilde{J}_0}{\partial k_m} - \frac{\partial \widetilde{J}_0}{\partial k_q}\Big) = (k_m - k_q)(\widetilde{J}'(k_m) - \widetilde{J}'(k_q)) \overset{(a)}{=} &(k_m - k_q)(J'(k_m) - J'(k_q)),\\ &\forall\ k_m, k_q \in \{1, 2, ..., p\}.\end{aligned} \tag{100}$$

where (a) holds by (70) and (b) is due to (99). When evaluated around $k_m \in \{1, 2, ..., p\}$ and $k_q = 0$, the resultant partial derivative reads

$$\begin{aligned}(k_m - k_q)\Big(\frac{\partial \widetilde{J}_0}{\partial k_m} - \frac{\partial \widetilde{J}_0}{\partial k_q}\Big) = k_m(\widetilde{J}'(k_m) - \widetilde{J}'(0)) &\overset{(a)}{=} k_m(J'(k_m) - \widetilde{J}'(0))\\ &\overset{(b)}{=} k_m(J'(k_m) - J'(1)) \overset{(c)}{=} \geq 0,\end{aligned} \tag{101}$$

where (a) holds by (70), (b) follows from (71), and (c) is true due to (99). Clearly, the function $\widetilde{J}_0$ is permutation invariant, together with (100) and (101) ensures that $\widetilde{J}_0$ satisfies the conditions of Lemma 4, and thus is Schur-convex. □