Amortized Inference for Causal Structure Learning

Abstract

Learning causal structure poses a combinatorial search problem that typically involves evaluating structures using a score or independence test. The resulting search is costly, and designing suitable scores or tests that capture prior knowledge is difficult. In this work, we propose to *amortize the process of causal structure learning*. Rather than searching over causal structures directly, we train a variational inference model to predict the causal structure from observational/interventional data. Our inference model acquires domain-specific inductive bias for causal discovery solely from data generated by a simulator. This allows us to bypass both the search over graphs and the hand-engineering of suitable score functions. Moreover, the architecture of our inference model is permutation invariant w.r.t. the data points and permutation equivariant w.r.t. the variables, facilitating generalization to significantly larger problem instances than seen during training. On synthetic data and semi-synthetic gene expression data, our models exhibit robust generalization capabilities under substantial distribution shift and significantly outperform existing algorithms, especially in the challenging genomics domain.

1 Introduction

A fundamental task in various scientific disciplines is to learn the causal structure among the variables we observe (Spirtes et al., 2000; Pearl, 2009). However, inferring this causal structure from observations of the variables is a difficult inverse problem. The solution space of potential causal structures, usually modeled as directed graphs, grows superexponentially with the number of variables. To infer a causal structure, standard methods have to search over potential graphs, usually maximizing either a graph scoring function or testing for conditional independencies (Heinze-Deml et al., 2018).

A key limitation of search approaches to causal discovery is the difficulty of specifying realistic inductive biases. On the one hand, score-based methods use strong assumptions like linearity (Shimizu et al., 2006), specific noise models (Hoyer et al., 2008; Peters and Bühlmann, 2014), or the absence of measurement error (cf. Scheines and Ramsey 2016; Zhang et al. 2017a), which are difficult to verify (Dawid, 2010; Reisach et al., 2021). On the other hand, constraint-based methods do not have enough domain-specific inductive bias, since even with oracle conditional independence tests, they are limited to identifying equivalence classes that may be exponentially large (He et al., 2015). Moreover, the search over directed graphs itself introduces unwanted bias and artifacts (Colombo et al., 2014). Its intractability imposes hard constraints on the causal structure, e.g., the node degree (Spirtes et al., 2000; Koivisto and Sood, 2004), limiting its suitability in real-world domains.

In the present work, we *amortize* the process of causal structure learning. In other words, our goal is to optimize an inference model to solve the inverse problem of inferring a causal structure from a provided data set. Our key insight is the dichotomy that, given a causal model, it is easy to generate realistic data—the inverse direction, i.e., inferring the causal structure from data, is hard. Many high-impact domains of causal discovery can be realistically forward-simulated, for example, the chemical ODEs and technical noise of gene regulatory networks (Schaffter et al., 2011; Dibaeinia

and Sinha, 2020), fMRI brain responses (Buxton, 2009; Bassett and Sporns, 2017), or chemical
kinetics (Anderson and Kurtz, 2011; Wilkinson, 2018). Leveraging such simulators for generating
training sets, we can use amortized inference to both acquire useful domain-specific inductive biases
not easily represented by a graph scoring function and bypass the problems of structure search.

The central challenge of amortizing causal discovery consists in designing a learning rule that
generalizes out-of-distribution. This generalization is essential because any simulated data
distribution is unlikely to cover all real-world distribution shifts. To quantify the epistemic uncertainty
in this inference process, we train an amortized variational inference model for causal discovery
(AVICI) by approximating the posterior over causal graphs given a data-generating distribution, e.g.,
a simulator. The model architecture of AVICI is permutation invariant and equivariant with respect
to the observation and variable dimensions, respectively, generalizing to significantly larger problem
instances than seen during training.

On synthetic data and semi-synthetic expression data of gene regulatory networks, our approach
significantly outperforms existing algorithms for causal discovery, often by a large margin. Moreover,
we demonstrate that our inference model induces calibrated uncertainties and robust behavior under
substantial distribution shifts of graphs, mechanisms, noise, and problem sizes, which makes our
trained models both reliable and versatile for future downstream use. In particular, AVICI was the
only one to infer plausible causal structure from realistic gene expression data, pushing the boundary
of what is possible in structure discovery of fields like molecular biology.

2 Background and Related Work

2.1 Causal Structure

In this work, we define a causal structure in the same way as Mooij et al. (2016). Other popular
models such as causal Bayesian networks and structural causal models (Pearl, 2009; Peters et al.,
2017) are less well-suited to describing systems with feedback loops, which we consider practically
relevant. We define the causal structure G of a directed graph whose nodes correspond to d variables
$\mathbf{x} = (x_1, \ldots, x_d)$ and edges represent all *direct causal* effects among the variables. A variable x_i has
a direct causal effect on x_j if intervening on x_i affects the outcome of x_j independent of the other
variables $\mathbf{x}_{\setminus ij} := \mathbf{x} \setminus \{x_i, x_j\}$, i.e., there exists $a \neq a'$ such that

$$p(x_j \mid \mathrm{do}(x_i = a, \mathbf{x}_{\setminus ij} = \mathbf{c})) \neq p(x_j \mid \mathrm{do}(x_i = a', \mathbf{x}_{\setminus ij} = \mathbf{c})) \tag{1}$$

for any $\mathbf{c}$. An *intervention* $\mathrm{do}(\cdot)$ denotes any active manipulation of the generative process of the
observations $\mathbf{x}$, like gene knockdowns in genomics, where the transcription rate of a gene is externally
reduced. G is an *abstraction* of the physical processes generating the data that describes the system's
behavior under arbitrary interventions (Peters et al., 2017; Rubenstein et al., 2017; Schölkopf, 2019).
We emphasize that G is not necessarily acyclic, however we later show how to apply our approach
when G is constrained to be acyclic. In this work, we assume causal sufficiency, i.e., that there are no
unobserved variables beyond $\mathbf{x}$ that are causal parents of more than one x_i (Peters et al., 2017).

2.2 Related Work

Classical methods search over causal graphs and evaluate them using a likelihood or conditional
independence tests (Chickering, 2003; Kalisch and Bühlman, 2007; Hauser and Bühlmann, 2012;
Zheng et al., 2018; Heinze-Deml et al., 2018). Others combine constraint and score-based ideas
(Tsamardinos et al., 2006) or use noise properties of the SCM (Shimizu et al., 2006; Hoyer et al.,
2008). In addition, there is work on leveraging multi-domain data (Tian and Pearl, 2001; Tillman
and Eberhardt, 2014; Ghassami et al., 2018; Zhang et al., 2017b; Huang et al., 2019, 2020). However,
these multi-domain approaches assume a shared causal model across datasets and thus do not study
the problem of generalizing to unseen causal systems.

Deep learning has been applied to causal inference, e.g., for the estimation of treatment effects
(Shalit et al., 2017; Louizos et al., 2017; Yoon et al., 2018) and in instrumental variable analysis
(Hartford et al., 2017; Bennett et al., 2019). In causal structure learning, neural networks have mostly
been used to model nonlinear causal mechanisms among the variables (Goudet et al., 2018; Yu
et al., 2019; Lachapelle et al., 2020; Brouillard et al., 2020; Ke et al., 2020; Lorch et al., 2021).
While amortized inference has gained increasing popularity in recent years (Amos, 2022), prior work

applying it to causal discovery only studied narrowly defined subproblems like the bivariate case or fixed causal mechanisms (Lopez-Paz et al., 2015; Löwe et al., 2022). In parallel work, Ke et al. (2022) independently perform causal discovery as a supervised learning task. There are significant differences in their approach, in particular, our inference models generalizes to significantly larger problem sizes, even when trained on realistic genomics data.

3 AVICI: Amortized Variational Inference for Causal Discovery

3.1 The Variational Objective

To amortize the process of causal structure learning, we define a data-generating distribution $p(D)$ that models the domain over which structure learning is performed. From a generative perspective, the observations $D = \{\mathbf{x}^1, \ldots, \mathbf{x}^n\} \sim p(D)$ are produced by sampling from a distribution over causal structures $p(G)$ and then obtaining realizations of a data-generating mechanism $p(D \mid G)$. It is important to highlight that $p(D \mid G)$ does not need to represent an ancestral sampling procedure over a directed acyclic graph. In particular, real-world systems are often more naturally modelled at different granularities or as dynamical systems (Mooij et al., 2013; Hoel et al., 2013; Schölkopf, 2019). Rather, $p(D \mid G)$ characterizes a data-generating process whose causal graph G formalizes all direct causal effects (Eq. 1) of the system's mechanisms—how intervening on certain variables influences the realizations of others (Mooij et al., 2016; Peters et al., 2017; Rubenstein et al., 2017).

Having observed a set of observations D in our domain, our goal is to approximate the true posterior over causal structures $p(G \mid D)$ with a variational distribution $q(G; \theta)$. To amortize this inference task for the data-generating distribution $p(D)$, we optimize an inference model f_ϕ to predict the variational parameters θ that minimize the *forward KL divergence* from the intractable posterior $p(G \mid D)$ for $D \sim p(D)$ (Ambrogioni et al., 2019):

$$\min_\phi \; \mathbb{E}_{p(D)} \, D_{KL}\big(p(G \mid D) \big\| q(G; f_\phi(D))\big) \tag{2}$$

Since it is not tractable to compute the true posterior in (2), we make use of ideas by Barber and Agakov (2004) and rewrite the expected forward KL to arrive at an equivalent, tractable objective.

$$\mathbb{E}_{p(D)} \, D_{KL}\big(p(G \mid D) \big\| q(G; f_\phi(D))\big) = \mathbb{E}_{p(D)} \, \mathbb{E}_{p(G \mid D)}[\log p(G \mid D) - \log q(G; f_\phi(D))]$$
$$= -\mathbb{E}_{p(G)} \, \mathbb{E}_{p(D \mid G)}[\log q(G; f_\phi(D))] + \text{const.} \tag{3}$$

The constant does not depend on ϕ, so we can maximize $\mathcal{L}(\phi) := \mathbb{E}_{p(G)} \, \mathbb{E}_{p(D \mid G)}[\log q(G; f_\phi(D))]$ to amortize variational inference for causal discovery (AVICI). While the data-generating distribution $p(D) = \sum_G p(D \mid G)p(G)$ can be arbitrarily complex, the training objective $\mathcal{L}$ is tractable whenever we have access to the ground truth causal graph G underlying the generative process of D, i.e., to samples from the joint distribution $p(G, D)$. In practice, $p(G)$ and $p(D \mid G)$ can thus be specified by use of a simulator for the domain of interest.

From an information-theoretic view, the objective (2) maximizes a variational lower bound on the mutual information $I[G; D]$ between the causal structure G and the observations D (Barber and Agakov, 2004). Starting from the definition of mutual information, we obtain

$$I[G; D] = H[G] - H[G \mid D] = H[G] + \mathbb{E}_{p(G,D)}[\log p(G \mid D)]$$
$$\geq H[G] + \mathbb{E}_{p(G,D)}[\log q(G; f_\phi(D))] = H[G] + \mathcal{L}(\phi) \tag{4}$$

where the entropy $H[G]$ is constant. The bound is tight if $\mathbb{E}_{p(D)} D_{KL}(p(G \mid D) \| q(G; f_\phi(D))) = 0$.

3.2 Likelihood-Free Inference using the Forward KL

The AVICI objective in (3) intentionally targets the forward KL $D_{KL}(p \| q)$, which requires optimizing $\mathbb{E}_{p(D \mid G)}[\log q(G; \theta)]$. This choice implies that we both model the density $q(G; \theta)$ explicitly and have access to samples from the true data-generating distribution. As long as *samples* from $p(G, D)$ are available for estimating $\mathcal{L}(\phi)$, minimizing the forward KL in (2) enables us to infer causal structures in arbitrarily complex domains—that is, even domains where it is difficult to specify an explicit functional form for $p(D \mid G)$. In addition, the forward KL typically yields more reliable uncertainty estimates since it does not suffer from the variance underestimation problems common for the backward KL (MacKay, 2003). Access to $p(G, D)$ is natural when a forward-model of the

Figure 1: Model architecture. (a) Our model maps the input to a three dimensional tensor of shape $n \times d \times k$ and remains permutation in- and equivariant over axes n and d, respectively. (b) Each layer self-attends first over axis n and then over d, sharing parameters across the other axis. (c) The inner-product of two variables' representations models the probability of a direct causal dependence.

domain exists, e.g., when sampling D synthetically from a structural equation model or a specialized domain simulator. We expect the quality of simulators in many domains to improve substantially over time, reducing the sim-to-real bias of the inference model f_ϕ.

In contrast, variational inference optimizes the reverse KL $D_{KL}(q\|p)$ with respect to θ, which involves the reconstruction term $\mathbb{E}_{q(G;\theta)}[\log p(D\,|\,G)]$ (Blei et al., 2017). This objective requires a parametric, tractable marginal likelihood $p(D\,|\,G)$. Unless jointly inferring the mechanism parameters (Brouillard et al., 2020; Lorch et al., 2021), this requirement limits inference to conjugate models with linear Gaussian or categorical mechanisms that assume zero measurement error (Geiger and Heckerman, 1994; Heckerman et al., 1995), which is not warranted in practice (Friston et al., 2000; Schaffter et al., 2011; Runge et al., 2019; Dibaeinia and Sinha, 2020). In addition, unless the noise scale is learned jointly, likelihoods can be sensitive to the measurement scale of $\mathbf{x}$ (Reisach et al., 2021).

In the following section, we describe a choice for the variational distribution $q(G;\theta)$ and the inference model f_ϕ that predicts θ given D. After that, we detail our training procedure of AVICI for optimizing the model parameters ϕ as well as for learning the inductive bias of acyclicity.

4 Inference Model

4.1 Variational Family

While any inference model that defines a density is feasible for maximizing the objective (3), we opt to use a factorized variational family in this work.

$$q(G;\theta) = \prod_{i,j} q(g_{i,j};\theta_{i,j}) \;\; \text{with} \;\; g_{i,j} \sim \text{Bern}(\theta_{i,j}) \tag{5}$$

The inference model f_ϕ thus maps a data set D corresponding to n samples $\{\boldsymbol{o}^1,\ldots,\boldsymbol{o}^n\}$ to a d-by-d matrix θ parameterizing the variational approximation of the causal graph. In addition to the joint observation $\mathbf{x}^i = (x_1^i,\ldots,x_d^i)$, each sample $\boldsymbol{o}^i = (o_1^i,\ldots,o_d^i)$ contains interventional information for each variable when available. For example, in the case of gene knockout interventions, we have $o_j^i = (x_j^i, u_j^i)$ and $u_j^i \in \{0,1\}$ indicating whether variable j was intervened upon in sample i. Other settings could be encoded analogously, e.g., when the intervention targets are unknown.

4.2 Model Architecture

To maximize sample efficiency, f_ϕ is constrained to satisfy the invariances inherent to the task of learning a causal structure from data. Firstly, f_ϕ should be *permutation invariant* across the sample dimension (axis n). Shuffling the samples should have no influence on the prediction, i.e., for any permutation π, $f_\phi(\pi(\{\boldsymbol{o}\})) = f_\phi(\{\boldsymbol{o}\})$. In addition, f_ϕ should be *permutation equivariant* across the variable dimension (axis d). Reordering the variables should permute the predicted causal edge probabilities, i.e., $f_\phi(\{\boldsymbol{o}_{\pi(1:d)}\})_{i,j} = f_\phi(\{\boldsymbol{o}_{1:d}\})_{\pi(i),\pi(j)}$. Lastly, f_ϕ should apply to any $d, n \geq 1$.

We parameterize f_ϕ as a deep neural network and introduce a novel architecture that satisfies the above invariants. After mapping each o_j^i to a real-valued vector using a position-wise linear layer, f_ϕ operates over a continuous, three dimensional tensor of n rows for the observations, d columns for the variables, and feature size k. Figure 1 illustrates the key components of the architecture.

Attending over axes n and d The core of our model is composed of $L = 8$ identical blocks. Each block consists of *four* residual layers, where the first and third apply multi-head self-attention and the

second and fourth position-wise feed-forward networks, similar to the vanilla Transformer encoder (Vaswani et al., 2017). To enable information flow across all $n \times d$ tokens of the representation, the model alternates in attending over the observation and the variable dimension. Specifically, the first self-attention block attends over axis n, treating axis d as a batch dimension; the second attends over axis d, treating axis n as a batch dimension. This design keeps memory requirements in check and results in a moderate model size because parameters are shared across the batched axes. Moreover, since multi-head attention is permutation equivariant and modules are shared across non-attended axes, the representation is permutation equivariant over axes n and d at all times (Lee et al., 2019b).

Causal graph model After building up a representation tensor from the input, we max-pool over the observation axis to obtain a representation $(\mathbf{z}^1, \ldots, \mathbf{z}^d)$ consisting of one vector $\mathbf{z}^i \in \mathbb{R}^k$ for each variable. Following Lorch et al. (2021), we use two linear layers to map each $\mathbf{z}^i$ into embeddings $\mathbf{u}^i, \mathbf{v}^i \in \mathbb{R}^k$ and finally model the probability of each edge in the causal graph as:

$$\theta_{i,j} = \sigma(\mathbf{u}^i \cdot \mathbf{v}^j) \tag{6}$$

where σ is the logistic function. Since max-pooling is invariant to permutations and since the graph model in (6) permutes with respect to axis d, the model satisfies the invariants of Section 4.2.

4.3 Acyclicity

There is no reason to generally exclude feedback loops in G. Cyclic causal effects often occur, e.g., when modeling stationary distributions of dynamical systems. Nonetheless, certain domains may be more accurately modeled by acyclic causal structures (Peters et al., 2017; Rubenstein et al., 2017). While the variational family in (5) does not allow enforcing it directly, we can optimize for acyclicity through ϕ. Whenever the acyclicity prior is justified, we amend the optimization problem in (2) with the constraint that q only models acyclic graphs in expectation:

$$\mathcal{F}(\phi) := \mathbb{E}_{p(D)} \left[h(f_\phi(D)) \right] = 0 \tag{7}$$

The function h is zero if and only if the predicted edge probabilities induce an acyclic graph. We use the insight by Lee et al. (2019a), who show that acyclicity is equivalent to the spectral radius ρ, i.e., the largest absolute eigenvalue, of the predicted matrix being zero. We use power iteration to approximate and differentiate through the largest eigenvalue of $f_\phi(D, I)$ (Golub and Van der Vorst, 2000; Lee et al., 2019a):

$$h(W) := \rho(W) \approx \frac{\mathbf{a}^\top W \mathbf{b}}{\mathbf{a}^\top \mathbf{b}} \quad \text{where for } t \text{ steps:} \quad \begin{array}{l} \mathbf{a} \leftarrow \mathbf{a}^\top W \ / \ \|\mathbf{a}^\top W\|_2 \\ \mathbf{b} \leftarrow W\mathbf{b} \ / \ \|W\mathbf{b}\|_2 \end{array} \tag{8}$$

and $\mathbf{a}, \mathbf{b} \in \mathbb{R}^d$ are randomly initialized. Since a few steps t are sufficient in practice, (8) scales with $O(d^2)$ and is significantly more efficient than alternative $O(d^3)$ constraints based on matrix powers (Zheng et al., 2018; Yu et al., 2019). We do not backpropagate gradients with respect to ϕ through $\mathbf{a}, \mathbf{b}$.

4.4 Optimization

Combining the objective in (3) with our inference model (5), we can directly use stochastic optimization to train the parameters ϕ of our neural network model. The expectations over $p(G, D)$ inside $\mathcal{L}$ and $\mathcal{F}$ are approximated using samples from the data-generating process of the domain, e.g, a simulator. In practice, the available data can be further augmented

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Algorithm 1 Training the AVICI inference model f_ϕ

Parameters: ϕ variational, λ dual, η step size
while not converged **do**
    for l steps **do**
        $\Delta\phi \propto \nabla_\phi \big(\mathcal{L}(\phi) - \lambda\mathcal{F}(\phi) \big)$
    $\lambda \leftarrow \lambda + \eta\mathcal{F}(\phi)$

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using bootstrapping. In case we impose an acyclicity prior on G, we solve the constrained program $\max_\phi \mathcal{L}(\phi)$ s.t. $\mathcal{F}(\phi) = 0$ through its dual formulation (Nandwani et al., 2019):

$$\min_\lambda \max_\phi \ \mathcal{L}(\phi) - \lambda\mathcal{F}(\phi) \tag{9}$$

Algorithm 1 summarizes the general optimization procedure for q_ϕ, which converges to a local optimum under regularity conditions on the learning rates (Jin et al., 2020). Without an acyclicity constraint, training reduces to the primal updates of ϕ with $\lambda = 0$.

Figure 2: Moving out-of-distribution in the RFF domain. Randomly sampled data-generating components of the RFF domain during training $p(D)$ (top) and o.o.d. evaluation $\tilde{p}(D)$ (bottom). For visualization, the adjacency matrices (a) are topologically sorted, the mechanisms (b) have two causal parents, where c and ℓ are output and length scales of the GP, and the noise (c) is shown as a function of one parent.

5 Experimental Setup

Evaluating causal discovery algorithms is difficult since there are there few interesting real-world data sets that come with ground-truth causal structure. Often the believed ground truths may be incomplete or change as expert knowledge improves (Schaffter et al., 2011). Following prior work, we deal with this issue by using simulated data with known causal structure.

5.1 Domains and Simulated Components

Our goal is to evaluate our inference model's ability to learn inductive biases in previously intractable domains and to generalize robustly beyond the data distribution $p(D)$ seen during training. We consider three domains: two classes of structural causal models (SCMs) as well as semi-synthetic single-cell expression data of gene regulatory networks (GRNs). To study generalization, we carefully construct a test distribution $\tilde{p}(D)$ that incurs substantial shift from $p(D)$ in terms of the underlying causal structures, mechanisms, and noise components.

Data-generating processes $p(D \mid G)$ The data-generating process of SCMs is defined by recursively sampling $x_j \leftarrow f_j(\mathbf{x}_{\text{pa}(j)}, \epsilon_j)$ based on a set of mechanisms f_j and endogenous noise distributions $p_j(\epsilon_j)$ for each variable. We consider the two domains of linear functions (LINEAR) and nonlinear functions of random Fourier features (RFF), which correspond to samples from a Gaussian process with squared exponential kernel (Rahimi and Recht, 2007). For the out-of-distribution (o.o.d.) setting $\tilde{p}(D)$, we draw the linear function weights and the kernel scales from the tails of $p(D)$ or even unseen value ranges. While using additive homoscedastic Gaussian noise for $p(D)$, we employ heteroscedastic Cauchy or Laplacian noise in $\tilde{p}(D)$ with $f_j(\mathbf{x}_{\text{Pa}(j)}, \epsilon_j) = f_j(\mathbf{x}_{\text{pa}(j)}) + h_j(\mathbf{x}_{\text{pa}(j)})\epsilon_j$ for some random function h_j. Interventions are performed by fixing selected variables x_j to a random value during sampling.

In addition to SCMs, we consider the challenging domain of GRNs (GRN) using the simulator by Dibaeinia and Sinha (2020). Contrary to SCMs, single-cell samples of gene expressions correspond to draws from the steady-state of a stochastic dynamical system that varies across cell types (Huynh-Thu and Sanguinetti, 2019). Like LINEAR and RFF, the simulator parameters considered o.o.d. are drawn from significantly wider ranges than for $p(D)$. The simulator additionally models technical noise effects, resulting in highly zero-inflated matrices D. When moving o.o.d., we use noise levels of different single-cell RNA sequencing technologies as calibrated by (Dibaeinia and Sinha, 2020). In GRN, interventions correspond to gene knockouts and are performed by forcing the expression rate of a variable x_j to zero.

Whenever a method infers structure from interventional data in our experiments, half of the dataset consists of observational data and half of single-variable interventions. In LINEAR and RFF, these single-variable interventions are performed on a random subset containing 50% of the nodes. Since GRN is more challenging, the interventional data here contains single-node interventions on all variables.

Causal structures $p(G)$ Following prior work, we use random graph models to generate ground truth causal structures. In all three domains, the training data distribution $p(D)$ is induced by

simple Erdős-Rényi and scale-free causal structures of various degrees (Erdős and Rényi, 1959; Barabási and Albert, 1999). In the o.o.d. setting, the data distributions $\tilde{p}(D)$ of the LINEAR and RFF domains are simulated using causal structures from different graph models: the Watts-Strogatz model, capturing small-world phenomena (Watts and Strogatz, 1998), the stochastic block model, generalizing Erdős-Rényi to community structures (Holland et al., 1983), and geometric random graphs, modeling connectivity based on spatial distance (Gilbert, 1961). In the GRN domain, we use subgraphs of the known *S. cerevisiae* and *E. coli* GRNs and the effect signs whenever known. To extract these subgraphs, we use the procedure by Marbach et al. (2009) to maintain structural patterns like motifs and modularity (Ravasz et al., 2002; Shen-Orr et al., 2002).

To illustrate the shift from $p(D)$ to $\tilde{p}(D)$, Figure 2 shows a set of graph, mechanism, and noise distribution samples in the RFF domain. Details on the simulator by Dibaeinia and Sinha (2020) and subgraph extraction are given in Appendix A. The specific parameter configurations and functions defining $p(D)$ and all $\tilde{p}(D)$ in the three domains are provided in Appendix B.

5.2 Evaluation Metrics

To asses how well the the predicted causal structure reflects the ground truth graph, we report the structural hamming distance (SHD) as well as the structural intervention distance (SID) (Peters and Bühlmann, 2015). While the SHD simply reflects the graph edit distance, the SID quantifies the closeness of two graphs in terms of their interventional adjustment sets. In addition, we follow existing work and evaluate the predicted uncertainty estimates $q(g_{i,j}; \theta_{i,j})$ over the set of possible edges by computing the areas under the precision-recall curve (AUPRC) and the receiver operating characteristic (AUROC) (Friedman and Koller, 2003; Lorch et al., 2021). For comparison with classical methods, we also collapse these uncertainties to hard predictions using a decision threshold of 0.5 and compute precision, recall, and the F1 score. More details are given in Appendix B.

5.3 Inference Model Configuration

We train three inference models overall, one for each domain, and perform all experiments on these three trained models. Performance could be further improved, e.g., by training models for more narrow settings specializing them solely on inference from observational data. However, to demonstrate the versatility of AVICI, we use the same models both for evaluation with only observational data as well as combinations of observations and interventional data. All inference models share identical hyperparameters for the architecture and optimization except the dropout rate. In held-out tasks of RFF and GRN, we found that dropout does not hurt performance in-distribution, so we increased the dropout rates from 0.0 to 0.1 and 0.3, respectively, to help o.o.d. generalisation. We add the acyclicity constraint for the SCM domains LINEAR and RFF. The data sets from $p(D)$ during training have $d = 5$ to 50 variables and $n = 200$ samples, where with probability 0.5, the observations contain 50 interventional samples. Details on the optimization and architecture are given in Appendix C.

6 Experimental Results

6.1 Out-Of-Distribution Generalization

We study the generalization capabilities of our inference models across the spectrum of test distributions described in Section 5.1. This analysis consists of two parts. First, we consider the model's sensitivity to queries with o.o.d. data. Second, we investigate how well it generalizes to different problem sizes.

Sensitivity to distribution shift We consider causal discovery from $n = 1000$ observations of systems of $d = 30$ variables. Starting from the training distribution $p(D)$, we incrementally introduce the described o.o.d. shifts in the causal structures, causal mechanisms, and finally noise, where fully o.o.d. corresponds to $\tilde{p}(D)$. The top row of Figure 3 visualizes the results of an empirical o.o.d. sensitivity analysis. The radar plots disentangle how combinations of the three o.o.d. aspects, i.e., graphs, mechanisms, and noise, affect the empirical performance of the AVICI models differently on the three domains LINEAR, RFF, and GRN. In addition to the metrics in Section 5.2, we also report the percentage of predicted graphs that are acyclic. To make the structural distances comparable, we use their ratios $\text{SID}_{\text{ratio}} := \text{SID}_{\text{in-dist.}}/\text{SID}$ and $\text{SHD}_{\text{ratio}}$, respectively. Given this, all metrics are in a similar value range and higher values indicate better performance.

(a) LINEAR   **(b)** RFF   **(c)** GRN

Figure 3: Generalization properties of the inference model f_ϕ**.** Top row shows performance metrics of AVICI under increasing distributional shift given $n = 1000$ observations for $d = 30$ variables. $\text{SID}_{\text{ratio}}$ is defined as $\text{SID}_{\text{in-dist.}}/\text{SID}$ and analogously for SHD. Bottom row shows the in-distribution AUPRC for various d as we vary the number of observations provided to AVICI. All metrics are the mean over fifteen random task instances and AVICI predicts from mixed observational and interventional data. For all metrics, higher is better. Error bars indicate the interquartile range.

As LINEAR is the easiest domain, AVICI performs very well across all metrics and hardly suffers from performance degradation due to distribution shift. In contrast, GRN is the most challenging, with performance degrading more significantly far o.o.d.. We observe that AVICI sometimes performs better o.o.d. than in-distribution, e.g., in GRN. This is because AVICI is empirically better at predicting edges adjacent to large-degree nodes, a predominant feature of the GRN graph distribution extracted from *E. coli* and *S. cerevisiae* that is not present in Erdős-Rényi graphs of the training distribution. The acyclicity constraint is perfectly satisfied for LINEAR and RFF.

We observe that AUPRC and AUROC do not suffer as much from distributional shift as the metrics based on thresholded point estimates. This indicates that the marginal posterior edge probabilities convey useful information about how certain AVICI is about individual edge predictions and how likely it is that the thresholded prediction is wrong. Overall, AVICI is surprisingly robust, even under substantial distribution shifts of all data distribution aspects at the same time.

Generalization to different problem sizes Next, we investigate the ability of our inference models to generalize to unseen problem sizes in terms of number of variables d and number of data points n. To this end, we test the models on increasingly larger sets of variables and various numbers of observations per task instance, here, with in-distribution data.

The bottom row of Figure 3 illustrates the AUPRC for AVICI's edge prediction when varying d and n. The predictions improve with the number of data points while exhibiting diminishing marginal improvement when seeing additional data. Moreover, the performance decreases smoothly as the number of variables increases and the task becomes harder. Most importantly, this robust behavior can be observed well beyond the settings ($n = 200$ and $d \le 50$) used during training.

6.2 Benchmarking

Finally, we compare AVICI with several causal discovery algorithms using either purely observational data or a mix of both observational and interventional data. The benchmark is performed on the fully o.o.d. domain distributions $\tilde{p}(D)$, i.e., under distribution shifts on causal graphs, mechanisms,

Table 1: Benchmark Results ($d = 30$ variables). Mean performance of all methods on 30 random task instances. Top section reports metrics for observational-only methods, the bottom section for methods using using observational and interventional data. The standard error is given in parentheses.

Algorithm	LINEAR		RFF		GRN	
	SID	F1	SID	F1	SID	F1
GES	458.2 (52.9)	0.038 (0.01)	458.2 (52.9)	0.038 (0.01)	**623.2 (30.9)**	0.031 (0.01)
LiNGAM	413.4 (48.4)	0.369 (0.04)	410.3 (47.6)	0.238 (0.02)	**606.2 (33.7)**	0.053 (0.01)
PC	412.1 (53.9)	0.325 (0.02)	370.5 (51.9)	0.441 (0.03)	**623.2 (33.6)**	0.038 (0.01)
DAG-GNN	474.3 (50.9)	0.156 (0.01)	425.0 (50.3)	0.230 (0.03)	**595.8 (35.2)**	**0.070 (0.01)**
AVICI (ours)[†]	**151.4 (19.8)**	**0.648 (0.04)**	**255.4 (48.9)**	**0.609 (0.06)**	630.1 (34.6)	0.026 (0.02)
GIES	**119.2 (26.4)**	0.734 (0.03)	304.5 (44.0)	0.337 (0.04)	597.8 (33.7)	0.074 (0.01)
IGSP	244.0 (34.4)	0.559 (0.02)	374.1 (45.0)	0.407 (0.04)	**495.1 (32.6)**	0.076 (0.01)
DCDI	384.8 (45.0)	0.318 (0.03)	286.8 (45.7)	0.412 (0.04)	614.4 (31.3)	0.065 (0.02)
AVICI (ours)	**117.7 (21.5)**	**0.808 (0.02)**	**196.6 (43.1)**	**0.701 (0.06)**	454.0 (45.9)	**0.262 (0.06)**

[†] Only observational data

and noise distributions. For each class of methods, we consider at least one constraint-based and one score-based approach as well as one recent algorithm using neural networks for modeling nonlinear mechanisms. In the observational setting, we compare AVICI with the PC algorithm (Spirtes et al., 2000), greedy equivalence search (GES) (Chickering, 2003), LiNGAM (Shimizu et al., 2006), and DAG-GNN (Yu et al., 2019). When using interventional data, we compare with greedy interventional equivalence search (GIES) (Hauser and Bühlmann, 2012), IGSP (Wang et al., 2017), and DCDI (Brouillard et al., 2020). For each domain, we independently calibrate the hyperparameters of each baseline in a grid search on held-out task instances. When computing the evaluation metrics, we favor methods that only predict (interventional) Markov equivalence classes by orienting undirected edges correctly when present in the ground truth. Details on the baselines are given in Appendix D.

Table 1 shows the SID and F1 scores of all methods for causal discovery from $n = 1000$ observations in systems of $d = 30$ variables. The AVICI model trained on LINEAR outperforms all baselines, both when given only observational data and with interventional data, despite operating with significant distribution shift. Only GIES achieves comparable accuracy. The same holds for RFF, where constraint-based and nonlinear approaches outperform GES and GIES as expected, but ultimately do not reach the accuracy set by AVICI. Finally, for GRN, we find that the inductive bias encoded by classical methods reaches its limits. However, provided interventional data, AVICI is able to use its learned inductive bias to infer plausible causal structure from the noisy gene expression data, which is a promising step towards reliable structure discovery in fields like molecular biology. Extended results for $d = 100$ variables are given in Appendix E. Overall, the methods' relative performance remains unchanged except for LINEAR, where GIES slightly outperforms AVICI.

7 Discussion

We proposed AVICI, a method for inferring causal structure by performing amortized variational inference over an arbitrary data-generating distribution. Our approach leverages the insight that inductive biases crucial for statistical efficiency might more easily be encoded in a domain simulator than in an inference technique. This is reflected in our experiments, where AVICI solves structure learning problems on complex domains intractable for existing approaches (Dibaeinia and Sinha, 2020). Our method can likely be extended to other typically difficult domains, including settings where we cannot assume causal sufficiency (Bhattacharya et al., 2021).

Using AVICI comes with several trade-offs that a practitioner must consider. First, while our experiments showed that optimizing the dual program helps induce acyclicity, this constraint cannot be satisfied with certainty. Moreover, similar to most amortized inference techniques (Amos, 2022), AVICI provides no theoretical guarantees of performance, whereas classical methods can in the infinite sample limit (Peters et al., 2017). Further work might achieve similar guarantees for AVICI as previously obtained in the bivariate setting (Lopez-Paz et al., 2015). Finally, while our trained models are fast at inference time, they can be intensive to train. To mitigate this training cost and make our results accessible to other researchers, we release the models studied in this work. Our experiments demonstrate that the trained models are highly robust to simulator misspecification, suggesting that these models could be useful even in causal structure learning tasks outside the domains studied in this paper.

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