Differentially Private Online-to-Batch for Smooth Losses

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Abstract

We develop a new reduction that converts any online convex optimization algorithm suffering $O(\sqrt{T})$ regret into an ϵ-differentially private stochastic convex optimization algorithm with the optimal convergence rate $\tilde{O}(1/\sqrt{T} + \sqrt{d}/\epsilon T)$ on smooth losses in linear time, forming a direct analogy to the classical non-private "online-to-batch" conversion. By applying our techniques to more advanced adaptive online algorithms, we produce adaptive differentially private counterparts whose convergence rates depend on apriori unknown variances or parameter norms.

1 Introduction

Solving stochastic convex optimization (SCO) problems forms a core component of training models in machine learning, and is the topic of this paper. The SCO problem is to optimize an objective $\mathcal{L}$:

$$\min_{x \in W} \mathcal{L}(x) = \mathop{\mathbb{E}}_{z \sim P_z} [\ell(x, z)] \tag{1}$$

Here, x represents model parameters or weights lying in a convex domain $W \subset \mathbb{R}^d$, P_z is some unknown distribution over examples z and $\ell(x, z)$ represents a loss function we will assume to be convex and smooth in x. Although we do not know P_z, we do know the loss ℓ, and we have access to an i.i.d. dataset $Z = (z_1, \ldots, z_T)$ that may have been obtained via user surveys or volunteer tests. Using this information, we would like to solve the optimization problem (1). The quality of a putative solution $\hat{x}$ is measured by the *suboptimality* gap $\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star)$ for $x_\star \in \operatorname{argmin}_{x \in W} \mathcal{L}(x)$.

In addition to solving (1), we also wish to *preserve privacy* for the people who contributed to the dataset Z. To this end, we require our algorithms to be *differentially private* (Dwork et al., 2006; Dwork and Roth, 2014), which means that replacing any one z_t with a different z_t' has a negligible effect on $\hat{x}$. There is a delicate interplay between privacy, dataset size, and solution quality. As T grows, the influence of any individual z_t on $\hat{x}$ decreases, increasing privacy. However, for any finite T, one must necessarily leak *some* information about the dataset in order to achieve a nontrivial $\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star)$. The goal, then, is to minimize $\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star)$ subject to the privacy constraint.

This problem has been well-studied in the literature, and by now the optimal tradeoffs are known and achievable[1]. In particular, Bassily et al. (2019) exhibits a polytime algorithm that achieves

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[1]We focus on the harder stochastic optimization problem rather than empirical risk minimization, which is also well-studied, e.g. (Chaudhuri et al., 2011; Kifer et al., 2012).

$\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star) \leq \tilde{O}(1/\sqrt{T} + \sqrt{d}/\epsilon T)$ where ϵ is a measure of privacy loss (large ϵ means less private), and moreover they show that this bound is tight in the worst case. Then, Feldman et al. (2020) provides an improved algorithm that obtains the same guarantee in $O(T)$ time. Further work on this problem considers assumptions on the geometry (Asi et al., 2021), gradient distributions (Kamath et al., 2021), or smoothness (Bassily et al., 2020; Kulkarni et al., 2021).

All private optimization algorithms we are aware of fall into one of two camps: either they employ some simple pre-processing that "sanitizes" the inputs to a non-private optimization algorithm (e.g. the empirically popular DP-SGD of Abadi et al. (2016)), or they make a careful analysis of the dynamics of their algorithm (e.g. bounding the sensitivity of a single step of stochastic gradient descent). In the former case, the algorithm typically does *not* achieve the optimal convergence rate for stochastic optimization. In the latter case, the algorithm becomes more rigidly tied to the privacy analysis, resulting in delicate "theory-crafted" methods that are less popular in practice.

In contrast, in the non-private setting, there is a simple and general technique to produce stochastic optimization algorithms with optimal convergence guarantees: the *online-to-batch conversion* (Cesa-Bianchi et al., 2004). This method directly converts any *online convex optimization* (OCO)(Shalev-Shwartz, 2011; Hazan, 2019; Orabona, 2019) algorithm into a stochastic optimization algorithm. OCO is a simple and elegant game-theoretic formulation of the optimization problem that has witnessed an explosion of diverse algorithms and techniques, so that this conversion result immediately implies a vast array of practical optimization algorithms. In summary: producing *private* stochastic optimization algorithms with optimal convergence rates is currently delicate and difficult, while producing non-private algorithms is essentially trivial.

Our goal is to rectify this issue. To do so, we produce a new *differentially private* online-to-batch conversion. In direct analogy to the non-private conversion, our method converts any OCO algorithm into a *private* stochastic optimization algorithm. After using our conversion, any OCO algorithm that achieves the optimal regret (the standard measure of algorithm quality in OCO), will automatically achieve the optimal suboptimality gap of $\tilde{O}(1/\sqrt{T}+\sqrt{d}/\epsilon T)$. Our conversion has additional desirable properties: although convexity is required for the guarantee on suboptimality, it is *not* required for the privacy guarantee, meaning that the method can be easily applied to non-convex tasks (e.g. in deep learning). Further, by largely decoupling the privacy analysis from the algorithm design through this reduction, we can leverage the rich literature of OCO to build private algorithms with new *adaptive* guarantees. For two explicit examples, (1) we construct an algorithm guaranteeing $\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star) \leq \tilde{O}(\sigma/\sqrt{T} + \sqrt{d}/\epsilon T)$, where σ is the apriori unknown standard deviation in the gradient $\nabla \ell(w, z)$, and (2), we construct an algorithm guaranteeing $\mathcal{L}(\hat{x}) - \mathcal{L}(x_\star) \leq \tilde{O}(\|\hat{x} - x_1\|/\sqrt{T} + \sqrt{d}/\epsilon T)$, where x_1 is any user-supplied point. This last guarantee may have applications in private *fine tuning* (Li et al., 2022; Yu et al., 2021; Hoory et al., 2021; Kurakin et al., 2022; Mehta et al., 2022), as the bound automatically improves when the algorithm is provided with a good initialization point.

1.1 Preliminaries

Problem setup We define the loss function as $\ell : W \times Z \to \mathbb{R}$ where $W \subseteq \mathbb{R}^d$ is a convex domain and $Z = (z_1, \ldots, z_T)$ is an element of $\mathcal{Z}^T$ for some dataspace $\mathcal{Z}$. We assume Z is an i.i.d. dataset and $z_t \sim P_z$ for some unknown disribution over $\mathcal{Z}$. We then define $\mathcal{L}(x) = \mathbb{E}_{z \sim P_z}[\ell(x, z)]$.

Let $\|\cdot\|$ be a norm on $\mathbb{R}^d$, with dual norm $\|\cdot\|_*$ defined by $\|g\|_* = \sup_{\|x\| \leq 1} \langle g, x \rangle$. By definition, $\langle g, x \rangle \leq \|g\|_* \|x\|$, (*Fenchel-Young's inequality*). We make the following assumptions:

Assumption 1. $\|\cdot\|^2$ is λ-strongly convex w.r.t. $\|\cdot\|$: for all $x, y \in \mathbb{R}^d$ and $g \in \partial\|x\|^2$,

$$\|y\|^2 \geq \|x\|^2 + \langle g, y - x \rangle + \frac{\lambda}{2}\|y - x\|^2.$$

Assumption 2. W has diameter at most D: $\forall\, x, y \in W, \|x - y\| \leq D$.

Assumption 3. $\ell(x, z)$ is G-Lipschitz in x: $\forall\, x \in W, z \in Z, \|\nabla\ell(x, z)\|_* \leq G$.

Assumption 4. $\ell(x, z)$ is H-smooth in x: $\forall\, x, y \in W, z \in Z, \|\nabla\ell(x, z) - \nabla\ell(y, z)\|_* \leq H\|x - y\|$.

Assumption 5. $\mathbb{E}[\|\nabla\mathcal{L}(x) - \nabla\ell(x, z)\|_*^2] \leq \sigma_G^2$ for all x, z.

Assumption 6. $\mathbb{E}[\|[\nabla\mathcal{L}(x) - \nabla\mathcal{L}(y)] - [\nabla\ell(x, z) - \nabla\ell(y, z)]\|_*^2] \leq \sigma_H^2\|x - y\|^2$ for all x, y, z.

Note that if ℓ is G-Lipschitz and H-smooth in x, then so it $\mathcal{L}$, and Assumption 5 and 6 hold with $\sigma_G \leq 2G$ and $\sigma_H \leq 2H$. Moreover, notice that Assumption 2 - 6 depend on the norm $\|\cdot\|$.

Differential Privacy We now provide a formal definition of our privacy metric, differential privacy (DP). The definition hinges on the notion of *neighboring datasets*: datasets $Z = (z_1, \ldots, z_T)$ and $Z' = (z_1', \ldots, z_T')$ in $\mathcal{Z}^T$ are said to be neighbors if $|Z - Z'| \triangleq |\{t \mid z_t \neq z_t'\}| = 1$.

Definition 1 ((ϵ, δ)-DP (Dwork and Roth, 2014)). A randomized algorithm $M : \mathcal{Z}^T \rightarrow \mathbb{R}^d$ is (ϵ, δ)-differentially private for $\epsilon, \delta \geq 0$ if for any neighboring $Z, Z' \in \mathcal{Z}^T$ and $S \in \mathbb{R}^d$:

$$\mathcal{P}\{M(Z) \in S\} \leq \exp(\epsilon)\, \mathcal{P}\{M(Z') \in S\} + \delta$$

An alternative definition is *Rényi differential privacy* (RDP), which is a generalization of DP that allows us to compose mechanisms more easily and achieve tighter privacy bounds in certain cases.

Definition 2 ((α, ϵ)-RDP (Mironov, 2017)). A randomized mechanism $M : \mathcal{Z}^T \rightarrow \mathbb{R}^d$ is said to be (α, ϵ)-Rényi differentially private for $\alpha > 1, \epsilon \geq 0$ if for any neighboring datasets $Z, Z' \in \mathcal{Z}^T$:

$$D_\alpha(M(Z)\|M(Z')) \leq \epsilon$$

where $D_\alpha(P\|Q) \triangleq \frac{1}{\alpha-1} \log \mathbb{E}_{x \sim Q}\left(\frac{P(x)}{Q(x)}\right)$.

RDP can be easily converted to the usual (ϵ, δ)-DP as follows (Mironov, 2017): if a randomized algorithm M is (α, ϵ)-RDP, then it is also ($\epsilon + \frac{\log 1/\delta}{\alpha-1}, \delta$)-DP for all $\delta \in (0, 1)$. In particular, if M is ($\alpha, \alpha\rho^2/2$)-RDP for all $\alpha > 1$, then it is also ($2\rho\sqrt{\log(1/\delta)}, \delta$)-DP for all $\delta \geq \exp(-\rho^2)$.

Throughout the paper, we also frequently use the notion of sensitivity:

Definition 3 (Sensitivity). The sensitivity of a function $f : \mathcal{Z}^T \rightarrow \mathbb{R}^d$ w.r.t. norm $\|\cdot\|$ is:

$$\Delta_f = \sup_{|Z-Z'|=1} \|f(Z) - f(Z')\|_*.$$

Almost all methods for ensuring differential privacy involve injecting noise whose scale increases with the sensitivity of the output. In other words, small sensitivity implies high privacy.

2 Diffentially Private Online-to-Batch

In this section, we present our main differentially private online-to-batch algorithm. To start, we need to define *online convex optimization* (OCO). OCO is a "game" in which for T rounds, $t = 1, \ldots, T$, an algorithms predicts a parameter $w_t \in W$. It then receives a convex loss $\ell_t : W \rightarrow \mathbb{R}$ and pays the loss $\ell_t(w_t)$. The quality of the algorithm is measured by the regret w.r.t. a competitor u, defined as $\text{Regret}_T(u) = \sum_{t=1}^T \ell_t(w_t) - \ell_t(u)$.

Online-to-batch algorithms (Cesa-Bianchi et al., 2004) convert OCO algorithms (online learners) into stochastic convex optimization algorithms. In particular, for $\beta_1, \ldots, \beta_T > 0$, the anytime online-to-batch conversion (Cutkosky, 2019) defines the t-th loss as $\ell_t(w) = \langle \beta_t \nabla \ell(x_t, z_t), w \rangle$, where $x_t = \sum_{i=1}^t \frac{\beta_i w_i}{\beta_{1:t}}$.[2] Convergence of anytime online-to-batch builds on the following key result, and its proof is in Appendix A for completeness.

Theorem 1 (Cutkosky (2019)). *For any sequence of $\beta_t > 0, g_t \in \mathbb{R}^d$, suppose an online learner predicts w_t and receives t-th loss $\ell_t(w) = \langle g_t, w \rangle$. Define $x_t = \sum_{i=1}^t \frac{\beta_i w_i}{\beta_{1:t}}$ where $\beta_{1:t} = \sum_{i=1}^t \beta_i$. Then for any convex and differentiable $\mathcal{L}$,*

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(u)) \leq \text{Regret}_T(u) + \sum_{t=1}^T \langle \beta_t \nabla \mathcal{L}(x_t) - g_t, w_t - u \rangle, \ \forall u \in \mathbb{R}^d.$$

A tighter bound is possible (Joulani et al., 2020), but the simpler expression above suffices for our analysis. As an immediate result, choosing $\beta_t = 1$ and $g_t = \beta_t \nabla \ell(x_t, z_t)$ satisfies $\mathbb{E}[g_t|x_t] = \beta_t \nabla \mathcal{L}(x_t)$, so $\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(u)] \leq \mathbb{E}[\text{Regret}_T(u)]/T$. Therefore, any online learner with sub-linear regret guarantees convergence for the last iterate x_T. Notice that due to the formulation of the anytime online-to-batch, the iterate x_t is *stable*. Consider the case where $\beta_t = 1$ for all t. Then, we have $\|x_t - x_{t-1}\| = \|w_t - x_{t-1}\|/t \leq D/t$ *regardless of what the online learner does*. This guarantee is

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[2]Throughout this paper, we denote $\beta_{1:t} = \sum_{i=1}^t \beta_i$.

significantly stronger than the classical online-to-batch result of Cesa-Bianchi et al. (2004). We would like to take advantage of this stability to design our private online-to-batch algorithm. Intuitively, the algorithm has much lower sensitivity due to the stability of the iterates, which can be exploited to improve privacy.

Our goal will be to make the entire sequence $g_1, \ldots, g_T$ private, which, in turn, makes the entire algorithm private. To do so, we would like to add noise to the g_t's while still having g_t be a good estimate of $\nabla \mathcal{L}(x_t)$. In standard non-private online-to-batch, g_t is usually defined as the stochastic gradient $\nabla \ell(x_t, z_t)$. However, directly adding noise to $\nabla \ell(x_t, z_t)$ is not a good idea because its sensitivity is $O(1)$ (more specifically the sensitivity is bounded by $2G$ by Lipschitzness). Consequently, we need to add noise to g_t whose variance is of order $O(1/\epsilon)$ in order to achieve ϵ-differential privacy.

Instead, we express $g_t = \sum_{i=1}^t \delta_i$, where $\delta_i = \beta_i \nabla \ell(x_i, z_i) - \beta_{i-1} \nabla \ell(x_{i-1}, z_i)$ and $\beta_0 \equiv 0$. Since we assume ℓ is smooth, if we set $\beta_i = 1$ then $\|\delta_i\|_* \leq H\|x_i - x_{i-1}\| \leq DH/i$, i.e., the sensitivity of δ_i is $O(1/i)$. As a result, we can privately estimate g_t with error roughly $\tilde{O}(1/t\epsilon)$ using the *tree aggregation mechanism* (Dwork et al., 2010; Chan et al., 2011), an advanced technique that privately estimates running sums, such as our $\sum_{i=1}^t \delta_i$. Compared to directly adding noise to $\nabla \ell(x_t, z_t)$, this method adds less noise ($\tilde{O}(1/t\epsilon)$ compared to $O(1/\epsilon)$) and thus allows us to achieve the optimal rate.

On the other hand, after using these advanced aggregation techniques, g_t is no longer an conditionally unbiased estimator of $\beta_t \nabla \mathcal{L}(x_t)$. More specifically, although it remains the case that $\mathbb{E}_{z_i}[\delta_i | z_1, \ldots, z_{i-1}] = \beta_i \nabla \mathcal{L}(x_i) - \beta_{i-1} \nabla \mathcal{L}(x_{i-1})$, it is not necessarily true that $\mathbb{E}[g_t | z_1, \ldots, z_{t-1}] \neq \beta_t \nabla \mathcal{L}(x_t)$. Unbiasedness plays a key role in standard convergence analyses, but we will need a much more delicate analysis.

Moreover, although we mostly discuss the $\beta_t = 1$ case above for intuition, our algorithm is analyzed using the general case $\beta_t = t^k$ for $k \geq 1$. The guiding principle for this formula is the sensitivity of δ_t. For $k \geq 1$, the sensitivity of δ_t is of order $O(t^{k-1})$. For $k = 1$, this is a constant sensitivity, which is particularly intuitive for analysis. For $k = 0$ (i.e. the standard weighting in online-to-batch), the sensitivity is actually $O(1/t)$, which is much more complicated to analyze. In order to apply the tree aggregation easily, we want the sensitivity of δ_t to be polynomial in t, rather than the inverse polynomial $1/t$, so we define $\beta_t = t^k$ and ask $t \geq k$. Furthermore, in all cases except for the parameter-free case (Section 5), our results hold for all $k \geq 1$. In the parameter-free case, we choose $k = 3$ for algebraic reasons.

The pseudo-code is presented in Algorithm 1, which has linear time complexity. It is similar to anytime online-to-batch, while we replace g_t with the more complicated definition and add noise γ_t generated by the NOISE subroutine, which implements the tree aggregation. More specifically, given random noises $\{R_t\}$, NOISE(t) returns $\sum_{i \in I_t} R_i$, where I_t is the set of cumulative sums of the binary expansion of t. That is, for some $n \geq \lfloor \log_2(t) \rfloor + 1$, we define $\text{bin}(t) \in \{0,1\}^n$ by $t = \sum_{i=1}^n \text{bin}(t)[i]2^{n-i}$. Then, I_t consists of all non-zero sums of the form $\sum_{k=1}^i \text{bin}(t)[i]2^{n-i}$. For examples, $7 = 4 + 2 + 1$, so $I_7 = \{4, 6, 7\}$; $8 = 8$, so $I_8 = \{8\}$.

Algorithm 1 Differentially-Private Online-to-Batch

1: **Input:** OCO algorithm $\mathcal{A}$ with domain W, positive sequence $\{\beta_t\}$, distribution $\mathcal{D}$, dataset Z.
2: **Initialize:** Set $\beta_0 = 0, x_0 = 0$ and $g_0 = 0$. Set global variable $\mathcal{R} = \{\}$.
3: **for** $t = 1, \ldots, T$ **do**
4:      Get w_t from $\mathcal{A}$ and compute $x_t = (\beta_{1:t-1}x_{t-1} + \beta_t w_t)/\beta_{1:t}$.
5:      Compute gradient difference $\delta_t = \beta_t \nabla \ell(x_t, z_t) - \beta_{t-1} \nabla \ell(x_{t-1}, z_t)$.
6:      Update $g_t = g_{t-1} + \delta_t$ and generate noise $\gamma_t = \text{NOISE}(t)$.
7:      Send $\ell_t(w) = \langle g_t + \gamma_t, w \rangle$ to $\mathcal{A}$ as the t-th loss.

8: **function** NOISE(t)
9:      **Initialize:** Set $k = 0$, $I_t = \{\}$, $\text{bin}(t)$ be the binary encoding of t, and $n = \lfloor \log_2 t \rfloor + 1$.
10:      **for** $i = 1, \ldots, n$ **do**
11:          If $\text{bin}(t)[i] = 1$, update $k = k + 2^{n-i}$ and $I_t = I_t \cup \{k\}$.
12:      Generate noise $\tilde{R}_t \sim \mathcal{D}$, compute $R_t = \sigma_t \tilde{R}_t$, and update $\mathcal{R} = \mathcal{R} \cup \{R_t\}$.
13:      **Return** $\gamma_t = \sum_{i \in I_t} R_i$.

Convergence Following Theorem 1 and Fenchel-Young's inequality, Algorithm 1 satisfies:

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \le \text{Regret}_T(x^*) + \sum_{t=1}^{T}\langle\beta_t\nabla\mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^*\rangle \tag{2}$$

$$\le \text{Regret}_T(x^*) + \sum_{t=1}^{T}D\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_* + \sum_{t=1}^{T}D\|\gamma_t\|_*. \tag{3}$$

(3) decomposes the suboptimality gap $\mathcal{L}(x_T) - \mathcal{L}(x^*)$ into three components: (i) regret $\text{Regret}_T(x^*)$, (ii) error associated with variance of g_t, measured by $\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*$, and (iii) error from DP mechanism, measured by $\|\gamma_t\|_*$. To get a tight bound, we observe that $\beta_t\nabla\mathcal{L}(x_t) - g_t$ and γ_t are sums of conditionally mean-zero random vectors (Lemma 15 in Appendix A). With a martingale bound in high dimension with general norm (Lemma 13), we can derive the following bounds. In particular, we show (Lemma 15) that if Assumption 1 - 6 hold and set $\beta_t = t^k$, then

$$\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2] \le 4(k+1)^2(\sigma_G^2 + D^2\sigma_H^2)t^{2k-1}/\lambda, . \tag{4}$$

Moreover, we show (Lemma 16) that if $\mathbb{E}[R_t] = 0$ and $\mathbb{E}[\|R_t\|_*^2] \le \bar{\sigma}_t^2$, then

$$\mathbb{E}[\|\gamma_t\|_*^2] \le 2(\max_{i\le t}\bar{\sigma}_i^2)\log_2(2t)/\lambda. \tag{5}$$

Privacy The next step is to determine how much noise is sufficient for Algorithm 1 to be RDP. We first make the following assumption on the distribution $\mathcal{D}$.

Definition 4 ((V, α)-RDP distribution). *A distribution $\mathcal{D}$ on $\mathbb{R}^d$ is said to be a (V, α)-RDP distribution on $\|\cdot\|$ if it satisfies that for $R \sim \mathcal{D}$ (i) $\mathbb{E}[R] = 0$, (ii) $\mathbb{E}[\|R\|_*^2] \le V$, and (iii) for all $\rho > 0$ and $\mu, \mu' \in \mathbb{R}^d$, if $\sigma^2 \ge \|\mu - \mu'\|_*^2/\rho^2$, then $D_\alpha(\sigma R + \mu\|\sigma R + \mu') \le \alpha\rho^2/2$.*[3]

Remark 2. A standard (V, α)-RDP distribution on the L_2 norm is the multivariate normal distribution. Let $R \sim \mathcal{N}(0, I)$ then it is clear that $\mathbb{E}[R] = 0$ and $\mathbb{E}[\|R\|_2^2] = d$. For the third condition, we can show that for all μ, μ' and $\alpha > 1$, $D_\alpha(\mathcal{N}(\mu, \sigma^2 I)\|\mathcal{N}(\mu', \sigma^2 I)) \le \alpha\|\mu - \mu'\|_2^2/2\sigma^2$, which is further bounded by $\alpha\rho^2/2$ for all $\sigma^2 \ge \|\mu - \mu'\|_2^2/\rho^2$ (Lemma 18) . Therefore, $\mathcal{N}(0, I)$ is a (d, α)-RDP distribution for all $\alpha > 1$.

As its name suggests, adding noise sampled from RDP distribution is sufficient to make a deterministic algorithm RDP. Consider function $\hat{f}(Z) = f(Z) + \sigma R$, where $R \sim \mathcal{D}$ and $\mathcal{D}$ is a (V, α)-RDP distribution on $\|\cdot\|$. Let Δ be the sensitivity of f w.r.t $\|\cdot\|$. Set $\sigma^2 \ge \Delta^2/\rho^2$, then by Definition 4,

$$D_\alpha(\hat{f}(Z)\|\hat{f}(Z')) = D_\alpha(f(Z) + \sigma R\|f(Z') + \sigma R) \le \alpha\rho^2/2.$$

Thus, $\hat{f}$ is ($\alpha, \alpha\rho^2/2$)-RDP. Moreover, we can compose RDP mechanisms via the tree aggregation mechanism. Specifically, we set $\beta_t = t^k$ and define the variance σ_t^2 in Algorithm 1 as follows:

$$\sigma_t^2 = \frac{4(k+1)^2}{\rho^2}(G + H\max_{i\in[t]}\|w_i - x_{i-1}\|)^2\log_2(2T)t^{2k-2}, \tag{6}$$

We assume $\|w_i - x_{i-1}\| \le D$. Upon substituting (6) into (5), we get $\mathbb{E}[\|R_t\|_*^2] \le \bar{\sigma}_t^2 \le V\sigma_t^2$ and:

$$\mathbb{E}[\|\gamma_t\|_*^2] \le \frac{8(k+1)^2V(G+DH)^2}{\lambda\rho^2}\log_2^2(2T)t^{2k-2}. \tag{7}$$

The following theorem shows that Algorithm 1 is Rényi differentially private if we define σ_t^2 as in (6). Its proof is presented in Appendix B. Note that the privacy guarantee does *not* require i.i.d. Z.

Theorem 3. *Suppose $\|\cdot\|^2$ is λ-strongly convex, W is bounded by D, ℓ is G-Lipschitz and H-smooth, and $\mathcal{D}$ is a (V, α)-RDP distribution. If $\beta_t = t^k$ and σ_t^2 is as defined in (6), then Algorithm 1 is ($\alpha, \alpha\rho^2/2$)-DP for all datasets Z.*

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[3]Here we slightly abuse the notation. For random vectors X, Y, $D_\alpha(X\|Y)$ denotes the Rényi divergence of the underlying distributions of X and Y.

Main Result.   Now we can combine all the previous results to prove the privacy and convergence guarantee of our algorithm.

Theorem 4. *Suppose Assumption 1 - 6 hold, and $\mathcal{D}$ is a (V, α)-RDP distribution. If we set $\beta_t = t^k$ and define σ_t^2 as in (6), then Algorithm 1 is $(\alpha, \alpha\rho^2/2)$-RDP and $\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)]$ is bounded by:*

$$\frac{(k+1)\,\mathbb{E}[\text{Regret}_T(x^*)]}{T^{k+1}} + \frac{2(k+1)^2 D}{\sqrt{\lambda}}\left(\frac{\sigma_G + D\sigma_H}{\sqrt{T}} + \frac{\sqrt{2V}(G+DH)\log_2(2T)}{\rho T}\right).$$

Moreover, recall that the online learner receives t-th loss $\ell_t(w) = \langle g_t + \gamma_t, w\rangle$. It holds that

$$\mathbb{E}[\|g_t + \gamma_t\|_*^2] \le 3t^{2k}\left(G^2 + \frac{4(k+1)^2}{\lambda}\left(\frac{(\sigma_G + D\sigma_H)^2}{t} + \frac{2V(G+DH)^2\log_2^2(2T)}{\rho^2 t^2}\right)\right).$$

Remark 5.   As an example, let's consider the Gaussian distribution $\mathcal{N}(0, I)$ on the 2-norm, which is a (d, α)-RDP distribution for all $\alpha > 1$ (Remark 2). For many popular online learners (OSD, OMD, FTRL), if $\mathbb{E}[\|\nabla\ell_t(w_t)\|_*^2] \le \hat{G}^2$ for all t, then $\mathbb{E}[\text{Regret}_T(x^*)] \le O(D\hat{G}\sqrt{T})$. Hence, Theorem 4 with $\mathcal{D} = \mathcal{N}(0, I)$ implies that

$$\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)] = O\left(\frac{D(G + D\sigma_H)}{\sqrt{T}} + \frac{\sqrt{d}D(G+DH)\log T}{\rho T}\right).$$

This bound is of $\tilde{O}(1/\sqrt{T} + \sqrt{d}/\rho T)$, which can be translated to an equivalent $(\epsilon, \delta)-$DP bound of $\tilde{O}(1/\sqrt{T} + \sqrt{d\log(1/\delta)}/\epsilon T)$. This bound matches the optimal rate for private stochastic optimization with convex and smooth losses (Bassily et al. (2019)).

Proof of Theorem 4.   From Eq. (3), we have:

$$\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)] \le \frac{1}{\beta_{1:T}}\,\mathbb{E}\left[\text{Regret}_T(x^*) + D\sum_{t=1}^{T}(\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_* + \|\gamma_t\|_*)\right]$$

Recall the bounds of $\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2]$ and $\mathbb{E}[\|\gamma_t\|_*^2]$ in (4) and (7). By Jensen's inequality, $\mathbb{E}[\|X\|_*] \le \sqrt{\mathbb{E}[\|X\|_*^2]}$. Moreover, since $\beta_t = t^k$ and $k \ge 1$, it holds that $\beta_{1:T} \ge T^{k+1}/k+1$, so:

$$\le \frac{\mathbb{E}[\text{Regret}]}{\beta_{1:T}} + \frac{D}{\beta_{1:T}}\sum_{t=1}^{T}\frac{2(k+1)(\sigma_G + D\sigma_H)t^{k-\frac{1}{2}}}{\sqrt{\lambda}} + \frac{\sqrt{8V}(k+1)(G+DH)\log_2(2T)t^{k-1}}{\sqrt{\lambda}\rho}$$

$$\le \frac{(k+1)\,\mathbb{E}[\text{Regret}_T(x^*)]}{T^{k+1}} + \frac{2(k+1)^2 D}{\sqrt{\lambda}}\left(\frac{\sigma_G + D\sigma_H}{\sqrt{T}} + \frac{\sqrt{2V}(G+DH)\log_2(2T)}{\rho T}\right).$$

For the second part of the theorem,

$$\mathbb{E}[\|g_t + \gamma_t\|_*^2] \le 3\,\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t)\|_*^2 + \|g_t - \beta_t\nabla\mathcal{L}(x_t)\|_*^2 + \|\gamma_t\|_*^2]$$

We bound the first term by Lipschitzness, the second by (4), and the third by (7)

$$\le 3t^{2k}G^2 + \frac{12(k+1)^2(\sigma_G^2 + D^2\sigma_H^2)t^{2k-1}}{\lambda} + \frac{24V(k+1)^2(G+DH)^2\log_2^2(2T)t^{2k-2}}{\lambda\rho^2}$$

$$= 3t^{2k}\left(G^2 + \frac{4(k+1)^2}{\lambda}\left(\frac{(\sigma_G + D\sigma_H)^2}{t} + \frac{2V(G+DH)^2\log_2^2(2T)}{\rho^2 t^2}\right)\right).$$

$\square$

3   The Optimistic Case

In this section, we show that choosing an optimistic online learner (Chiang et al., 2012; Rakhlin and Sridharan, 2013; Steinhardt and Liang, 2014) will accelerate our DP online-to-batch algorithm. Optimistic algorithms are provided with additional "hints" in the form of $\hat{\ell}_t(w) = \langle\hat{g}_t, w\rangle$ as an

approximation of the true loss $\ell_t(w) = \langle g_t, w \rangle$, and they can incorporate $\hat{\ell}_t$ to decide w_t. The regret of an optimistic algorithm depends on the quality of hints: if $\hat{g}_t \approx g_t$, then it achieves a low regret. Formally, in this paper, we say an online learning algorithm is *optimistic w.r.t. norm* $\|\cdot\|$ if its regret is the following:

$$\text{Regret}_T(x^*) \leq O\left(D\sqrt{\sum_{t=1}^{T} \|\hat{g}_t - g_t\|_*^2}\right), \tag{8}$$

where D denotes the diameter of the learner's domain.

A common choice of the hint $\hat{g}_t$ is g_{t-1}, the gradient in the last round since intuitively, one could expect $g_{t-1} \approx g_t$ when the loss functions are smooth. In this section, we also follow this choice. Recall that in Algorithm 1, the online learner receives t-th loss $\ell_t(w) = \langle g_t + \gamma_t, w \rangle$, where $g_t = \sum_{i=1}^{t} \delta_i$ is the sum of gradient differences, and γ_t is some noise. Therefore, we define the t-th hint as $\hat{g}_t = g_{t-1} + \gamma_{t-1}$.

Theorem 6. *Suppose Assumption 1 - 4 hold, and $\mathcal{D}$ is a (V, α)-RDP distribution. Set $\beta_t = t^k$ and σ_t^2 as defined in (6), If the online learner is optimistic (satisfying (8)) with t-th gradient $\bar{g}_t = g_t + \gamma_t$ and t-th hint $\hat{g}_t = \bar{g}_{t-1}$, then*

$$\frac{\mathbb{E}[\text{Regret}_T(x^*)]}{\beta_{1:T}} \leq O\left(D(G + DH)\left(1 + \frac{\sqrt{V}\log T}{\sqrt{\lambda}\rho}\right)\frac{1}{T^{3/2}}\right).$$

The proof is in Appendix D. As an immediate corollary, if we further assume Assumption 5 and 6, then Theorem 4 applies. Together with this theorem, they imply that optimistic learners achieve the following convergence rate that is adaptive to the variance:

$$\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)] = O\left(\frac{D(\sigma_G + D\sigma_H)}{\sqrt{T}} + \frac{\sqrt{d}D(G + DH)\log T}{\rho T}\right).$$

Compared to the bound in the non-optimistic case (Remark 5), this bound has σ_G instead of G in the first term. Thus, when the gradient $\nabla \ell(x_t, z_t)$ has low variance, i.e., $\sigma_G \ll G$, the optimistic bound outperforms the standard bound.

4 The Strongly Convex Case

In this section, we prove that in the case of strong convexity, our algorithm can be improved by regularizing the loss of online learner. If $\mathcal{L}$ is strongly convex, then we can prove a similar result to Theorem 1 (the proof is in Appendix E).

Lemma 7. *Suppose $\mathcal{L}$ is μ-strongly convex w.r.t. $\|\cdot\|$. If we replace $\ell_t(w) = \langle g_t + \gamma_t, w \rangle$ with $\bar{\ell}_t(w) = \ell_t(w) + \frac{\beta_t \mu}{4}\|w - x_t\|^2$ in Algorithm 1, and denote the associated regret by $\overline{\text{Regret}}_T$, then*

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^{T} \langle \beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^* \rangle - \frac{\beta_t \mu}{8}\|w_t - x^*\|^2$$

$$\leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^{T} \frac{2\|\beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t\|_*^2}{\beta_t \mu}.$$

Compared to Lemma 1 and Equation 3, in the strongly convex case, there is an additional term $-\frac{\beta_t \mu}{8}\|w_t - x^*\|^2$, which allows the improved convergence rate:

Theorem 8. *Suppose Assumption 1 - 6 hold, and $\mathcal{D}$ is a (V, α)-RDP distribution. Also suppose $\mathcal{L}$ is μ-strongly convex. Set $\beta_t = t^k$ and σ_t^2 as defined in (6). Then $\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)]$ is bounded by:*

$$\frac{(k+1)\mathbb{E}[\overline{\text{Regret}}_T(x^*)]}{T^{k+1}} + \frac{16(k+1)^3}{\lambda\mu}\left(\frac{(\sigma_G + D\sigma_H)^2}{T} + \frac{2V(G + DH)^2 \log_2^2(2T)}{\rho^2 T^2}\right).$$

Moreover, for all $\bar{g}_t \in \partial\bar{\ell}_t(w_t)$,

$$\mathbb{E}[\|\bar{g}_t\|_*^2] \leq t^{2k}\left(4G^2 + \mu^2 D^2 + \frac{16(k+1)^2}{\lambda}\left(\frac{(\sigma_G + D\sigma_H)^2}{t} + \frac{2V(G + DH)^2 \log_2^2(2T)}{\rho^2 t^2}\right)\right).$$

Remark 9. As an example, again consider the case $\mathcal{D} = \mathcal{N}(0, I)$ with L_2 norm. Online subgradient descent (OSD) with appropriate learning rate on μ_t-strongly convex losses ℓ_t achieves:

$$\text{Regret}_T(u) \leq \sum_{t=1}^{T} \frac{\|g_t\|_2^2}{2\sum_{i=1}^{t}\mu_i},$$

where $g_t \in \partial\ell_t(w_t)$. In our case, $\|\cdot\|_2^2$ is 2-strongly convex and the regularized loss $\bar{\ell}_t$ is $\frac{\beta_t\mu}{2}$-strongly convex, i.e. $\mu_t = \mu t^k/2$ and $\sum_{i=1}^{t}\mu_i = O(\mu t^{k+1})$. Therefore, by Theorem 8,

$$\mathbb{E}[\overline{\text{Regret}}_T(x^*)] \leq O\left(\frac{T^k}{\mu}\left(G^2 + \mu^2 D^2 + \frac{(\sigma_G + D\sigma_H)^2}{T} + \frac{d(G + DH)^2\log^2 T}{\rho^2 T^2}\right)\right).$$

Consequently,

$$\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)] \leq O\left(\frac{(G + \mu D + D\sigma_H)^2}{\mu T} + \frac{d(G + DH)^2\log^2 T}{\mu\rho^2 T^2}\right).$$

This again matches the optimal private convergence rates.

Proof of Theorem 8. We have already bounded $\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2]$ and $\mathbb{E}[\|\gamma_t\|_*^2]$ in (4) and (7) respectively, so

$$\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t - \gamma_t\|_*^2] \leq 2\,\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2 + \|\gamma_t\|_*^2]$$
$$\leq \frac{8(k+1)^2 t^{2k-1}}{\lambda}\left((\sigma_G + D\sigma_H)^2 + \frac{2V(G + DH)^2\log_2^2(2T)}{\rho^2 t}\right).$$

Upon substituting this into Lemma 7 and replace $\beta_t = t^k$, we get:

$$\beta_{1:T}\,\mathbb{E}[\mathcal{L}(x_T) - \mathcal{L}(x^*)]$$
$$\leq \mathbb{E}[\overline{\text{Regret}}_T(x^*)] + \sum_{t=1}^{T}\frac{16(k+1)^2 t^{k-1}}{\lambda\mu}\left((\sigma_G + D\sigma_H)^2 + \frac{2V(G + DH)^2\log_2^2(2T)}{\rho^2 t}\right)$$
$$\leq \mathbb{E}[\overline{\text{Regret}}_T(x^*)] + \frac{16(k+1)^2}{\lambda\mu}\left((\sigma_G + D\sigma_H)^2 T^k + \frac{2V(G + DH)^2\log_2^2(2T)T^{k-1}}{\rho^2}\right).$$

Dividing both sides by $\beta_{1:T} \geq T^{k+1}/(k+1)$ proves the first part of the theorem.

For the second part, recall that $\bar{\ell}_t(w) = \langle g_t + \gamma_t, w\rangle + \frac{\beta_t\mu}{4}\|w - x_t\|^2$. Therefore, for all $\bar{g}_t \in \partial\bar{\ell}_t(w)$, $\bar{g}_t = g_t + \gamma_t + \frac{\beta_t\mu}{4}v$, where $v \in \partial\|w_t - x_t\|^2$. We follow the same argument in Theorem 4:

$$\mathbb{E}[\|\bar{g}_t\|_*^2] \leq 4\,\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t)\|_*^2 + \|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2 + \|\gamma_t\|_*^2 + \|\tfrac{\beta_t\mu}{4}v\|_*^2]$$

Here $v \in \partial\|w_t - x_t\|^2$. We bound the first term by Lipschitz, the second by (4), and the third by (7). Moreover, by chain rule (Proposition 21), we can show that $\|v\|_* \leq 2D$, so:

$$\leq 4t^{2k}G^2 + \frac{16(k+1)^2(\sigma_G + D\sigma_H)^2 t^{2k-1}}{\lambda}$$
$$+ \frac{32V(k+1)^2(G + DH)^2 t^{2k-2}\log_2^2(2T)}{\lambda\rho^2} + \mu^2 D^2 t^{2k}$$
$$\leq t^{2k}\left(4G^2 + \mu^2 D^2 + \frac{16(k+1)^2}{\lambda}\left(\frac{(\sigma_G + D\sigma_H)^2}{t} + \frac{2V(G + DH)^2\log_2^2(2T)}{\rho^2 t^2}\right)\right).$$

$\square$

5 Parameter-free Algorithm

In this section, we apply Algorithm 1 with a "parameter-free" online learner. These are algorithms that guarantee $\text{Regret}_T(u) \leq \tilde{O}(\|u\|\sqrt{T})$ for all competitors u simultaneously (Orabona and Pál,

2016; Cutkosky and Orabona, 2018; Mhammedi and Koolen, 2020). By shifting coordinates, it is possible to obtain $\text{Regret}_T(u) \leq \tilde{O}(\|u - x_0\|\sqrt{T})$ for any pre-specified point x_0. Thus, if $x_0 \approx u$ is some good initialization, perhaps generated by pretraining, this bound yields significantly smaller regret than if we had used the worst-case diameter bound $\|u - x_0\| \leq D$.

In order to obtain this refined bound with privacy, we need to make a small modification to our conversion. For simplicity, we focus on Euclidean space with 2-norm, and we assume the distribution $\mathcal{D}$ is in addition sub-Gaussian, i.e., for $R \sim \mathcal{D}$,

$$\mathcal{P}\{\sup_{\|a\|_2=1} \langle R, a \rangle \geq \epsilon\} \leq \exp\left(-\frac{\epsilon^2}{2\sigma^2}\right).$$

In general, the proof extends to any Banach space and any distribution $\mathcal{D}$ that concentrates on it.

In previous analysis (Equation (3)), we roughly bound $\|w_t - x^*\| \leq D$. However, in this section, we come up with a finer high probability bound that maintains a dependence on $\|w_t\|, \|x^*\|$. We then replace the loss $\ell_t(w)$ in Algorithm 1 with a regularized loss $\ell_t(w) + \xi_t\|w\|_2 + \nu_t\|w\|_2^2$, and we show that the new algorithm with regularized loss can achieve a parameter-free bound. The complete proof is presented in Appendix F.

Theorem 10. *Suppose w.r.t. 2-norm, W is bounded by D and ℓ is G-Lipschitz and H-smooth. Suppose $\mathcal{D}$ is (V, α)-RDP distribution and is $\sigma_\mathcal{D}$-sub-Gaussian. If we set $\beta_t = t^3$ (i.e. $k = 3$) and set σ_t^2 as defined in (6), then with probability at least $1 - \delta$,*

$$\mathcal{L}(x_T) - \mathcal{L}(x^*) \leq \frac{4}{T^4}\left(\text{Regret}_T(x^*) + \sum_{t=1}^{T} \xi_t(\|w_t\|_2 + \|x^*\|_2) + \nu_t(\|w_t\|_2^2 + \|x^*\|_2^2)\right).$$

where C is a universal constant, $A = 8\sqrt{2}C^2$, $A' = 8\sqrt{d}\sigma_\mathcal{D}C^2$, $\kappa = 1 + DH/G$, and

$$\xi_t = AG\Phi t^{5/2} + A'(G + DH)\frac{\Phi \log_2(2T)t^2}{\rho}, \quad \nu_t = 28AH\Phi t^{5/2}, \quad \Phi = \sqrt{\log \frac{20dT\log(2\kappa T)}{\delta}}.$$

Theorem 11. *Following the assumptions and notations in Theorem 10, if we replace $\ell_t(w)$ in Algorithm 1 with regularized loss $\bar{\ell}_t(w) = \ell_t(w) + \xi_t\|w\|_2 + \nu_t\|w\|_2^2$ and denote the associated regret as $\overline{\text{Regret}}_t$, then with probability at least $1 - \delta$, $\mathcal{L}(x_T) - \mathcal{L}(x^*)$ is bounded by:*

$$\frac{4\overline{\text{Regret}}_T(x^*)}{T^4} + \frac{8A\|x^*\|(G + 28\|x^*\|H)\Phi}{\sqrt{T}} + \frac{8A'\|x^*\|(G + DH)\Phi \log_2(2T)}{\rho T}.$$

Moreover, with probability at least $1 - \delta$, for all t and for all $w \in W, \bar{g}_t \in \partial \bar{\ell}_t(w)$,

$$\|\bar{g}_t\|_2 \leq Gt^3 + A(2G + 57DH)\Phi t^{5/2} + 2A'(G + DH)\frac{\Phi \log_2(2T)t^2}{\rho}.$$

Remark 12. *If $\|\bar{g}_t\|_2 \leq \hat{G}$ for all t, parameter-free algorithms achieve regret bound $\text{Regret}_T(u) = \tilde{O}(\|u\|_2\hat{G}\sqrt{T})$. Therefore, with $\mathcal{D} = \mathcal{N}(0, I)$, $\mathcal{D}$ is 1-sub-Gaussian and $A' = O(\sqrt{d})$, so Theorem 11 implies that with probability $1 - \delta$,*

$$\frac{\overline{\text{Regret}}_T(x^*)}{T^4} = \tilde{O}\left(\frac{\|x^*\|_2 G}{\sqrt{T}} + \frac{\|x^*\|_2(G + DH)}{T} + \frac{\|x^*\|_2\sqrt{d}(G + DH)}{\rho T^{3/2}}\right).$$

Consequently,

$$\mathcal{L}(x_T) - \mathcal{L}(x^*) \leq \tilde{O}\left(\frac{\|x^*\|_2(G + DH)}{\sqrt{T}} + \frac{\|x^*\|_2\sqrt{d}(G + DH)}{\rho T}\right).$$

Proof of Theorem 11. The regularized regret satisfies

$$\overline{\text{Regret}}_T(x^*) = \text{Regret}_T(x^*) + \sum_{t=1}^{T} \xi_t(\|w_t\|_2 - \|x^*\|_2) + \nu_t(\|w_t\|_2^2 - \|x^*\|_2^2).$$

Hence, upon substituting this equation into Theorem 10, we get: with probability at least $1 - \delta$,

$$\mathcal{L}(x_T) - \mathcal{L}(x^*) \leq \frac{4\overline{\text{Regret}}_T(x^*)}{T^4} + \frac{8}{T^4} \sum_{t=1}^{T} \xi_t \|x^*\|_2 + \nu_t \|x^*\|_2^2$$

$$\leq \frac{4\text{Regret}_T(x^*)}{T^4} + \frac{8(AG\Phi\|x^*\|_2 + 28AH\Phi\|x^*\|_2^2)}{\sqrt{T}} + \frac{8A'(G+DH)\Phi \log_2(2T)\|x^*\|}{\rho T}.$$

The second inequality is from $\sum_{t=1}^{T} \xi_t \leq T\xi_T$ because ξ_t is increasing with t (so is ν_t).

For the second part of the theorem, for each fixed t and for all $\bar{g}_t \in \bar{\ell}_t(w_t)$, $\bar{g}_t = g_t + \gamma_t + \xi_t u + 2\nu_t w_t$, where $u \in \partial\|w_t\|_2$ and thus $\|u\|_2 \leq 1$. Therefore,

$$\|\bar{g}_t\|_2 \leq \|g_t - \beta_t \nabla \mathcal{L}(x_t)\|_2 + \|\beta_t \nabla \mathcal{L}(x_t)\|_2 + \|\gamma_t\|_2 + \|\xi_t u\|_2 + \|2\nu_t w\|_2$$

By Lipschitzness, $\|\beta_t \nabla \mathcal{L}(x_t)\|_2 \leq Gt^3$. Since W is bounded, $\|2\nu_t w_t\|_2 \leq 2D\nu_t$. Also, $\|\xi_t u\|_2 \leq \xi_t$. Moreover, we can prove (in Eq. 14 and 17) that for each t, with probability at least $1 - \delta/2T$,

$$\|\beta_t \nabla \mathcal{L}(x_t) - g_t\|_2 \leq 8C^2\Phi \sqrt{\sum_{i=1}^{t} i^4 (G + H\|w_i - x_{i-1}\|_2)^2} \leq A\Phi(G+DH)t^{5/2},$$

$$\|\gamma_t\|_2 \leq \frac{A'}{\rho}(G+DH)\Phi \log_2(2T)t^2.$$

Upon taking the union bound for all t and the definition of ξ_t, ν_t, we get the desired bound. $\qquad\square$

6 Conclusion

We have presented a new online-to-batch conversion that produces private stochastic optimization algorithms on smooth losses. Online algorithms achieving the optimal $O(\sqrt{T})$ regret automatically achieve the optimal $\tilde{O}(1/\sqrt{T} + \sqrt{d}/\epsilon T)$ convergence rate. Combining this technique with the literature on online learning can yield new private optimization algorithms.

Limitations: Our algorithm requires smoothness, and unlike some other bounds, we cannot tolerate large H. In the worst case when $H = \sqrt{T}$ and $\sigma_H = H$, our standard bound in Remark 5 becomes $O(1)$. In other words, we need to assume $H = o(\sqrt{T})$ to ensure a non-trivial bound. Removing this restriction would significantly improve the generality of the procedure,

The dependence on H comes from the sensitivity of δ_t (Lemma 15), where we apply smoothness to bound $\|\beta_t(\nabla\ell(x_t, z_t) - \nabla\ell(x_{t-1}, z_t))\| \leq \beta_t H\|x_t - x_{t-1}\|$ and use the stability of x_t to further bound $\|x_t - x_{t-1}\| \leq D\beta_t/\beta_{1:t}$, which are necessary steps in order to bound the sensitivity of δ_t by $O(t^{k-1})$. Hence, it's not clear how to remove the smoothness assumption.

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A Proofs for Convergence (Section 2)

Theorem 1 (Cutkosky (2019)). *For any sequence of $\beta_t > 0, g_t \in \mathbb{R}^d$, suppose an online learner predicts w_t and receives t-th loss $\ell_t(w) = \langle g_t, w \rangle$. Define $x_t = \sum_{i=1}^{t} \frac{\beta_i w_i}{\beta_{1:t}}$ where $\beta_{1:t} = \sum_{i=1}^{t} \beta_i$. Then for any convex and differentiable $\mathcal{L}$,*

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(u)) \leq \mathrm{Regret}_T(u) + \sum_{t=1}^{T} \langle \beta_t \nabla \mathcal{L}(x_t) - g_t, w_t - u \rangle, \; \forall u \in \mathbb{R}^d.$$

Proof. Since $\mathcal{L}$ is convex,

$$\sum_{t=1}^{T} \beta_t(\mathcal{L}(x_t) - \mathcal{L}(u)) \leq \sum_{t=1}^{T} \beta_t \langle \nabla \mathcal{L}(x_t), x_t - w_t + w_t - u \rangle$$

$$= \sum_{t=1}^{T} \langle \nabla \mathcal{L}(x_t), \beta_t(x_t - w_t) \rangle + \langle \beta_t \nabla \mathcal{L}(x_t) - g_t + g_t, w_t - u \rangle.$$

By construction of x_t, it holds that $\beta_t(x_t - w_t) = \beta_{1:t-1}(x_{t-1} - x_t)$. By convexity,

$$\langle \nabla \mathcal{L}(x_t), \beta_{1:t-1}(x_{t-1} - x_t) \rangle \leq \beta_{1:t-1} \left(\mathcal{L}(x_{t-1}) - \mathcal{L}(x_t) \right).$$

Next, we move $\sum \beta_t \mathcal{L}(x_t)$ to the right, giving:

$$-\beta_{1:T}\mathcal{L}(u) \leq \sum_{t=1}^{T} \left(\beta_{1:t-1}\mathcal{L}(x_{t-1}) - \beta_{1:t}\mathcal{L}(x_t) + \langle \beta_t \nabla \mathcal{L}(x_t) - g_t + g_t, w_t - u \rangle \right)$$

$$= -\beta_{1:T}\mathcal{L}(x_T) + \mathrm{Regret}_T(u) + \sum_{t=1}^{T} \langle \beta_t \nabla \mathcal{L}(x_T) - g_t, w_t - u \rangle,$$

where the equality follows from (i) the telescopic sum of $\beta_{1:t-1}\nabla \mathcal{L}(x_{t-1}) + \beta_{1:t}\nabla \mathcal{L}(x_t)$ and (ii) the definition of regret that $\mathrm{Regret}_T(u) = \sum_{t=1}^{T} \langle g_t, w_t - u \rangle$. $\qquad \square$

Lemma 13. *Suppose $\| \cdot \|^2$ is λ-strongly convex w.r.t. $\| \cdot \|$, and let $\{X_t\}$ be a sequence of random vectors such that (i) $\mathbb{E}[\|X_t\|_*] < \infty$ and (ii) $\mathbb{E}[X_{t+1} \mid X_{1:t}] = 0$ for all t. Then,*

$$\mathbb{E}\left[\left\| \sum_{t=1}^{T} X_t \right\|_*^2 \right] \leq \frac{2}{\lambda} \sum_{t=1}^{T} \mathbb{E}[\|X_t\|_*^2].$$

Proof. We use the regret approach to prove this statement. For simplicity, denote $M_T = \sum_{t=1}^{T} X_t$. Consider an online learner which receives $\ell_t(x) = \langle X_t, x \rangle$ as t-th loss and updates w_{t+1}. Then by definition of regret, for any u,

$$-\langle M_T, u \rangle \leq \mathrm{Regret}_T(u) - \sum_{t=1}^{T} \langle X_t, w_t \rangle.$$

Since w_t only depends on $X_{1:t-1}$ but *not* on X_t, w_t is constant given $X_{1:t-1}$. Therefore,

$$\mathbb{E}[\langle X_t, w_t \rangle] = \mathop{\mathbb{E}}_{X_{1:t-1}} \mathop{\mathbb{E}}_{X_t} \left[\langle X_t, w_t \rangle | X_{1:t-1} \right]$$

$$= \mathop{\mathbb{E}}_{X_{1:t-1}} \left[\langle \mathop{\mathbb{E}}_{X_t}[X_t | X_{1:t-1}], w_t \rangle \right] = 0,$$

where the second equality follows from the assumption that $\mathbb{E}[X_t | X_{1:t-1}] = 0$. Therefore,

$$\mathbb{E}[\langle M_T, -u \rangle] \leq \mathbb{E}[\mathrm{Regret}_T(u)].$$

Recall the definition of the dual norm that $\|M_T\|_* = \sup_{\|x\|=1} \langle M_T, x \rangle$. Therefore, if we define $u^* = \|M_T\|_* \mathrm{argmax}_{\|u\|=1} \langle M_T, -u \rangle$, then it holds that

$$\langle M_T, -u^* \rangle = \|M_T\|_*^2 \quad and \quad \|u^*\| = \|M_T\|_*.$$

Let the follow-the-regularized-leader (FTRL) algorithm be the online learner. Orabona (2019) proved that for any regularizer ψ that is λ-strongly convex w.r.t. $\|\cdot\|$, FTRL achieves the following regret:

$$\text{Regret}_T(u) \leq \frac{\psi(u)}{\eta} + \frac{\eta}{2\lambda}\sum_{t=1}^{T}\|X_t\|_*^2.$$

Since we assume $\|\cdot\|^2$ is λ-strongly convex w.r.t. $\|\cdot\|$, we can define $\psi(x) = \|x\|^2$ and get:

$$\mathbb{E}[\langle M_T, -u^*\rangle] = \mathbb{E}[\|M_T\|_*^2] \leq \mathbb{E}[\text{Regret}_T(u^*)] \leq \mathbb{E}\left[\frac{\|M_T\|_*^2}{\eta} + \frac{\eta}{2\lambda}\sum_{t=1}^{T}\|X_t\|_*^2\right].$$

Equivalently, upon moving terms around we have:

$$\mathbb{E}[\|M_T\|_*^2] \leq \mathbb{E}\left[\frac{\eta^2}{2\lambda(\eta-1)}\sum_{t=1}^{T}\|X_t\|_*^2\right] \leq \mathbb{E}\left[\frac{2}{\lambda}\sum_{t=1}^{T}\|X_t\|_*^2\right].$$

The second inequality holds because $\inf_\eta \frac{\eta^2}{\eta-1} = 4$ when $\eta = 2$. $\qquad\square$

In this paper, we will always set $\beta_t = t^k$ for some $k \geq 1$. The following proposition gives relevant bounds for $\beta_t - \beta_{t-1}$ and $\beta_t^2/\beta_{1:t}$.

Proposition 14. If $\beta_t = t^k$, then (i) $\beta_t - \beta_{t-1} \leq kt^{k-1}$ and (ii) $\beta_t^2/\beta_{1:t} \leq (k+1)t^{k-1}$.

Proof. For the first part, by mean value theorem, there exists some $\tau \in [t-1, t]$ such that

$$\beta_t - \beta_{t-1} = t^k - (t-1)^k = k\tau^{k-1} \leq kt^{k-1}.$$

For the second part, for any increasing function f, it holds that $\sum_{i=1}^{t} f(i) \geq \int_0^t f(x)\,dx$, so

$$\beta_{1:t} = \sum_{i=1}^{t} i^k \geq \int_0^t x^k\,dx = \frac{t^{k+1}}{k+1}.$$

Hence, $\beta_t^2/\beta_{1:t} \leq (k+1)t^{k-1}$. $\qquad\square$

Lemma 15. *Suppose $\|\cdot\|^2$ is λ-strongly convex, W is bounded by D, and ℓ is G-Lipschitz and H-smooth. If we set $\beta_t = t^k$, then*

$$\|\delta_t\|_* \leq (k+1)(G + H\|w_t - x_{t-1}\|)t^{k-1}.$$

If we further assume Assumption 5 and 6, then $\beta_t\nabla\mathcal{L}(x_t) - g_t = \sum_{i=1}^{t} X_i$ such that:

(i) $X_i = [\beta_i\nabla\mathcal{L}(x_i) - \beta_{i-1}\nabla\mathcal{L}(x_{i-1})] - [\beta_i\nabla\ell(x_i, z_i) - \beta_{i-1}\nabla\ell(x_{i-1}, z_i)]$,

(ii) $\mathbb{E}[X_i|z_{1:i-1}] = 0$, *and* (iii) $\mathbb{E}[\|X_i\|_*^2] \leq 2(k+1)^2(\sigma_G^2 + D^2\sigma_H^2)i^{2k-2}$.

Consequently,

$$\mathbb{E}[\|\beta_t\nabla\mathcal{L}(x_t) - g_t\|_*^2] \leq 4(k+1)^2(\sigma_G^2 + D^2\sigma_H^2)t^{2k-1}/\lambda.$$

Proof. For the first part, note that

$$\begin{aligned}\delta_t &= \beta_t\nabla\ell(x_t, z_t) - \beta_{t-1}\nabla\ell(x_{t-1}, z_t) \\ &= (\beta_t - \beta_{t-1})\nabla\ell(x_{t-1}, z_t) + \beta_t(\nabla\ell(x_t, z_t) - \nabla\ell(x_{t-1}, z_t)).\end{aligned}$$

Since ℓ is G-Lipschitz and H-smooth, $\|\nabla\ell(x_{t-1}, z_t)\|_* \leq G$ and

$$\|\nabla\ell(x_t, z_t) - \nabla\ell(x_{t-1}, z_t)\|_* \leq H\|x_t - x_{t-1}\| \leq (\beta_t/\beta_{1:t})H\|w_t - x_{t-1}\|,$$

where the second inequality follows from the definition of x_t that $\beta_{1:t}(x_t - x_{t-1}) = \beta_t(w_t - x_{t-1})$. The first result then follows from Proposition 14.

For the second part, by telescopic sum $\beta_t \nabla \mathcal{L}(x_t) = \sum_{i=1}^t \beta_i \nabla \mathcal{L}(x_i) - \beta_{i-1} \nabla \mathcal{L}(x_{i-1})$, and recall that $g_t = \sum_{i=1}^t \delta_i$. Therefore,

$$\beta_t \nabla \mathcal{L}(x_t) - g_t = \sum_{i=1}^t [\beta_i \nabla \mathcal{L}(x_i) - \beta_{i-1} \nabla \mathcal{L}(x_{i-1})] - [\beta_i \nabla \ell(x_i, z_i) - \beta_{i-1} \nabla \ell(x_{i-1}, z_i)].$$

We denote each summand by X_i, and we can check X_i satisfies condition (ii) and (iii). First, since we assume $\nabla \mathcal{L}(w) = \mathbb{E}_z[\nabla \ell(w, z)]$ for all w, it holds that $\mathbb{E}[X_i | z_{1:i-1}] = 0$. Second, we decompose X_i in the same way as the first part:

$$\mathbb{E}[\|X_i\|_*^2] = \mathbb{E}[\|(\beta_i - \beta_{i-1})[\nabla(\mathcal{L}(x_{i-1}) - \nabla \ell(x_{i-1}, z_i)] \\ + \beta_i([\nabla \mathcal{L}(x_i) - \nabla \ell(x_i, z_i)] - [\nabla \mathcal{L}(x_{i-1}) - \nabla \ell(x_{i-1}, z_i)])\|_*^2]$$

Recall the assumption of σ_G^2 and σ_H^2 and that $\|x_i - x_{i-1}\| \le \beta_i / \beta_{1:i} \|w_i - x_{i-1}\|$.

$$\le 2(\beta_i - \beta_{i-1})^2 \sigma_G^2 + 2\beta_i^2 \sigma_H^2 \|x_i - x_{i-1}\|^2$$
$$\le 2(\beta_i - \beta_{i-1})^2 \sigma_G^2 + 2(\beta_i^2 / \beta_{1:i})^2 \sigma_H^2 \|w_i - x_{i-1}\|^2$$
$$\le 2(k+1)^2 (\sigma_G^2 + D^2 \sigma_H^2) i^{2k-2}.$$

The last inequality follows from the assumption that $\|w_i - x_{i-1}\| \le D$ and Proposition 14.

The last part of the theorem is a direct result from Lemma 13:

$$\mathbb{E}[\|\nabla \mathcal{L}(x_t) - g_t\|_*^2] \le \frac{2}{\lambda} \sum_{i=1}^t \mathbb{E}[\|X_i\|_*^2] \le \frac{4(k+1)^2}{\lambda} (\sigma_G^2 + D^2 \sigma_H^2) t^{2k-1}.$$

$\square$

Lemma 16. *Suppose $\mathbb{E}[R_t] = 0$ and $\mathbb{E}[\|R_t\|_*^2] \le \bar{\sigma}_t^2$, then*

$$\mathbb{E}[\|\gamma_t\|_*^2] \le 2(\max_{i \le t} \bar{\sigma}_i^2) \log_2(2t)/\lambda.$$

Proof. By construction, $\gamma_t = \sum_{i \in I_t} R_i$, and R_i's are independent and mean-zero. Therefore, Lemma 13 can be applied, which yields

$$\mathbb{E}[\|\gamma_t\|_*^2] \le \frac{2}{\lambda} \sum_{i \in I_t} \mathbb{E}[\|R_i\|_*^2] \le \frac{2}{\lambda} \sum_{i \in I_t} \bar{\sigma}_i^2 \le 2(\max_{i \le t} \bar{\sigma}^2) \log_2(2t)/\lambda.$$

The last inequality is from the fact that $|I_t| \le \log_2(2t)$. $\square$

B Proofs for RDP (Section 2)

In this section, we prove the tree aggregation mechanism for RDP mechanisms implemented in Algorithm 1 correctly composes individual RDP mechanisms. Before that, we will first prove a general composition theorem for RDP.

B.1 Advanced Composition for RDP

Throughout this section, we use the subscript $1 : T$ to denote a sequence of T elements. We denote $\mathcal{Z}$ the data space and $Z = (z_1, \ldots, z_T), Z' = (z_1', \ldots, z_T')$ neighboring datasets in $\mathcal{Z}^T$ that differs only at the q-th element (i.e., $z_t \ne z_t'$ if and only if $t = q$). We consider RDP mechanisms $F_1, \ldots, F_T$ such that $F_1 : \mathcal{Z}^T \to W$ and $F_t : \mathcal{Z}^T \times W^{t-1} \to W$. We assume that for each t, there exists some index set $S_t \subseteq [T]$ such that F_t only depends on S_t. Formally, we assume:

Assumption 7. *Let $Z, Z' \in \mathcal{Z}$ be two neighboring datasets which differs only at the q-th element. Each F_t associates with $S_t \subseteq [T]$ such that if $q \notin S_t$, then $F_t(Z, x_{1:t-1}) = F_t(Z', x_{1:t-1})$ for all $x_{1:t-1} \in W^{t-1}$.*

For a fixed norm $\|\cdot\|$, we assume $\mathcal{D}$ is a (V, α)-RDP distribution on norm $\|\cdot\|$ (see Definition 4), and we define the sensitivity of F_t w.r.t. $\|\cdot\|$ as $\Delta_t(x_{1:t-1})$, a function of inputs $x_{1:t-1}$:

$$\Delta_t(x_{1:t-1}) = \sup_{|Z-Z'|=1} \|F_t(Z, x_{1:t-1}) - F_t(Z', x_{1:t-1})\|_*.$$

We also define the output as $\hat{f}_t = F_t(Z, \hat{f}_{1:t-1}) + \sigma_t \zeta_t$, where $\zeta_t \sim \mathcal{D}$ and $\sigma_t^2 \geq \Delta_t(\hat{f}_{1:t-1})^2/\rho^2$. In particular, σ_t only depends on $\hat{f}_{1:t-1}$ and does not depend on $\hat{f}_{t:T}$, i.e., the future. The pseudo-code of this composition is in Algorithm 2. For simplicity, we assume F's are deterministic mechanisms, but we can extend to random mechanisms by treating the random generator as part of the input.

Algorithm 2 Advanced Composition for RDP Mechanisms

1: **Input:** Dataset Z; functions $F_1, \ldots, F_T$ with sensitivity $\Delta_1, \ldots, \Delta_T$; (V, α)-RDP distribution $\mathcal{D}$; privacy constants $\rho_1, \ldots, \rho_T$.
2: Sample random $\zeta_1 \sim \mathcal{D}$ and compute $\sigma_1^2 \geq \Delta_1^2/\rho_1^2$.
3: Set $f_1 = F_1(Z)$ and $\hat{f}_1 = f_1 + \sigma_1 \zeta_1$
4: **for** $t = 2, \ldots, T$ **do**
5:     Sample random $\zeta_t \sim \mathcal{D}$ and compute $\sigma_t^2 \geq \Delta_t(\hat{f}_{1:t-1})^2/\rho_t^2$.
6:     Set $f_t = F_t(Z, \hat{f}_{1:t-1})$ and $\hat{f}_t = f_t + \sigma_t \zeta_t$.
7: **Return** $\hat{f}_1, \ldots, \hat{f}_T$.

By definition of RDP distribution, each $\hat{f}_t$ is $(\alpha, \alpha\rho^2/2)$-RDP. We will also show that the composition $(\hat{f}_1, \ldots, \hat{f}_T)$ is also RDP.

Theorem 17. *We define* $\text{IN}(q) = \{t : q \in S_t\}$ *and* $\text{OUT}(q) = \{t : q \notin S_t\}$. *If* $F_1, \ldots, F_T$ *satisfy Assumption 7, then Algorithm 2 is* (α, S)-*RDP, where*

$$S = \max_{q \in [T]} \sum_{t \in \text{IN}(q)} \alpha\rho_t^2/2.$$

As an immediate corollary, if we set $\rho_t = \rho$ for all t and define $U = \max_q |\text{IN}(q)|$, then Algorithm 2 is $(\alpha, U\alpha\rho^2/2)$-RDP.

Proof. Let Z, Z' be any neighboring datasets and assume they differ at q, and we denote $\hat{f}_t = F_t(Z, \hat{f}_{1:t-1}) + \sigma_t(\hat{f}_{1:t-1})\zeta_t$ and $\hat{f}'_t = F_t(Z', \hat{f}'_{1:t-1}) + \sigma_t(\hat{f}'_{1:t-1})\zeta_t$. In this proof, we use the notation $\sigma_t(\hat{f}_{1:t-1})$ to emphasize that σ_t satisfying $\sigma_t^2 \geq \Delta_t(\hat{f}_{1:t-1})^2/\rho_t^2$ is a function of $\hat{f}_{1:t-1}$.

The probability density of the joint distribution of $\hat{f}_{1:T}$, say P, and the density of $\hat{f}'_{1:T}$, say Q, are:

$$P(x_{1:T}) = \prod_{t=1}^{T} P_t(x_t|x_{1:t-1}), \quad Q(x_{1:T}) = \prod_{t=1}^{T} Q_t(x_t|x_{1:t-1}),$$

where $P_t(\cdot|x_{1:t-1})$ is the density of $(\hat{f}_t|\hat{f}_{1:t-1} = x_{1:t-1}) = F_t(Z, x_{1:t-1}) + \sigma_t(x_{1:t-1})\zeta_t$ and $Q_t(\cdot|x_{1:t-1})$ is the density of $F_t(Z', x_{1:t-1}) + \sigma_t(x_{1:t-1})\zeta_t$. Note that σ_t is the same for P_t and Q_t.

By definition of Rényi divergence,

$$D_\alpha(P\|Q) = \frac{1}{\alpha-1} \log \int P(x_{1:T})^\alpha Q(x_{1:T})^{1-\alpha} \, dx_{1:T}. \tag{9}$$

Previous multiplication rule implies that:

$$P(x_{1:T})^\alpha Q(x_{1:T})^{1-\alpha} = \prod_{t \in \text{IN}(q) \sqcup \text{OUT}(q)} \left(P_t(x_t|x_{1:t-1})^\alpha Q_t(x_t|x_{1:t-1})^{1-\alpha} \right)$$

For all $t \in \text{OUT}(q)$ (i.e., $q \notin S_t$), Assumption 7 implies that $F_t(Z, x_{1:t-1}) = F_t(Z', x_{1:t-1})$, so $P_t(\cdot|x_{1:t-1}) = Q_t(\cdot|x_{1:t-1})$ and

$$\int P_t(x_t|x_{1:t-1})^\alpha Q_t(x_t|x_{1:t-1})^{1-\alpha} \, dx_t = \int P_t(x_t|x_{1:t-1}) \, dx_t = 1. \tag{10}$$

The second inequality holds because $P_t(\cdot|x_{1:t-1})$ is a probability density.

On the other hand, for all $t \in \mathrm{IN}(q)$,

$$\sigma_t(x_{1:t-1})^2 \geq \Delta_t(x_{1:t-1})^2/\rho_t^2$$
$$\geq \|F_t(Z, x_{1:t-1}) - F_t(Z', x_{1:t-1})\|_*^2/\rho_t^2,$$

so by Definition 4,

$$D_\alpha(P_t\|Q_t) = \frac{1}{\alpha-1}\log \int P_t(x|x_{1:t-1})^\alpha Q_t(x|x_{1:t-1})^{1-\alpha}\,dx \leq \alpha\rho_t^2/2.$$

Equivalently,

$$\int P_t(x|x_{1:t-1})^\alpha Q_t(x|x_{1:t-1})^{1-\alpha}\,dx \leq \exp((\alpha-1)\alpha\rho_t^2/2). \tag{11}$$

Note that P_t, Q_t only depend on $x_{1:t-1}$ and *not* on $x_{t:T}$, so we can rearrange the integral in (9) as:

$$\int P(x_{1:T})^\alpha Q(x_{1:T})^{1-\alpha}\,dx_{1:T}$$
$$= \int P_T(x_T|x_{1:T-1})^\alpha Q_T(x_T|x_{1:T-1})^{1-\alpha}\cdots\left(\int P_1(x_1)^\alpha Q_1(x_1)^{1-\alpha}\,dx_1\right)\cdots dx_T$$

Evaluating the composite integral from inside to outside with (10) and (11) gives:

$$\leq \exp\left(\sum_{t\in\mathrm{IN}(q)}(\alpha-1)\alpha\rho_t^2/2\right).$$

In conclusion, for all $|Z - Z'| = 1$,

$$D_\alpha(P\|Q) = \frac{1}{\alpha-1}\log\int P(x_{1:T})^\alpha Q(x_{1:T})^{1-\alpha}\,dx_{1:T} \leq \max_q \sum_{t\in\mathrm{IN}(q)}\alpha\rho_t^2/2 = S.$$

$\square$

B.2 Algorithm 1 is RDP

Now we are ready to prove the tree aggregation in Algorithm 1.

Theorem 3. *Suppose $\|\cdot\|^2$ is λ-strongly convex, W is bounded by D, ℓ is G-Lipschitz and H-smooth, and $\mathcal{D}$ is a (V, α)-RDP distribution. If $\beta_t = t^k$ and σ_t^2 is as defined in (6), then Algorithm 1 is $(\alpha, \alpha\rho^2/2)$-DP for all datasets Z.*

Proof. Recall the definition of I_t in Algorithm 1: we define $s_0 = 0$, and $s_i = \max_k\{s_{i-1} + 2^k : 2^k | t - s_{i-1}\}$ until $s_n = t$ for some n, and we define $I_t = \{s_1, \ldots, s_n\}$. For example, $I_4 = \{4\}$ and $I_7 = \{4, 6, 7\}$. We then define $S_t = \{s_{n-1} + 1, s_{n-1} + 2, \ldots, t\}$ (e.g., $S_4 = \{1, 2, 3, 4\}$ and $S_7 = \{7\}$). Observe that $\{S_i : i \in I_t\}$ is a partition of $[t]$.

Let Z, Z' be neighboring datasets that differ at the q-th element. Define $F_t : \mathcal{Z}^T \times W^{t-1} \to W$ as:

$$F_t(Z, \hat{f}_{1:t-1}) = \sum_{i\in S_t}\delta_i(Z, \hat{f}_{1:t-1})$$
$$= \sum_{i\in S_t}\beta_i\nabla\ell(x_i, z_i) - \beta_{i-1}\nabla\ell(x_{i-1}, z_i),$$

where x_i's are the parameters as defined in Algorithm 1, and z_i is the i-th data in Z. We then define $\hat{f}_t = F_t(Z, \hat{f}_{1:t-1}) + \sigma_t(\hat{f}_{1:t-1})\tilde{R}_t$ and $\hat{f}'_t = F_t(Z', \hat{f}'_{1:t-1}) + \sigma_t(\hat{f}'_{1:t-1})\tilde{R}_t$, where $\tilde{R}_t \sim \mathcal{D}$.

For simplicity, we denote $\delta_t = \delta_t(Z, \hat{f}_{1:t-1})$ and $\sigma_t = \sigma_t(\hat{f}_{1:t-1})$ (and δ'_t, σ'_t respectively). We also denote x_i, w_i and x'_i, w'_i as parameters w.r.t. Z, Z' respectively. Since $\{S_i : i \in I_t\}$ partitions $[t]$,

$$g_t + \gamma_t = \sum_{i=1}^t \delta_i + \sum_{i\in I_t}R_i = \sum_{i\in I_t}\left(\sum_{j\in S_i}\delta_j + \sigma_i\tilde{R}_i\right) = \sum_{i\in I_t}\hat{f}_i.$$

By construction of Algorithm 1, $x_1, \ldots, x_t$ are determined by $\{g_i + \gamma_i\}_{i=1}^{t-1}$ and equivalently by $\hat{f}_{1:t-1}$. In particular, they do not depend on $f_{t:T}$. Consequently, this implies that (i) if $\hat{f}_{1:t-1} = \hat{f}'_{1:t-1}$ then $x_{1:t} = x'_{1:t}$ and $w_{1:t} = w'_{1:t}$ and (ii) if in addition $z_i = z'_i$ then $\delta_i = \delta'_i$.

This implies that F_t's satisfy Assumption 7: if $q \notin S_t$ (i.e., $z_i = z'_i$ for all $i \in S_t$), then for any fixed $f_{1:t-1} \in W^{t-1}$, $F_t(Z, f_{1:t-1}) = F_t(Z', f_{1:t-1})$ because $\delta_i = \delta'_i$ for all $i \in S_t$. Consequently, Theorem 17 can be applied, which states that if $\sigma_t(\hat{f}_{1:t-1})^2 \geq \Delta_t(\hat{f}_{1:t-1})^2/\rho^2$, then $(\hat{f}_1, \ldots, \hat{f}_T)$ is $(\alpha, U\alpha\rho^2/2)$-RDP where $U = \max_q |\text{IN}(q)|$ and $\text{IN}(q) = \{t : q \in S_t\}$. Note that $U \leq \log_2(2T)$.

The sensitivity of F_t at fixed $f_{1:t-1} \in W^{t-1}$ is bounded by:

$$\Delta_t(f_{1:t-1}) = \sup_{|Z-Z'|=1} \|F_t(Z, f_{1:t-1}) - F_t(Z', f_{1:t-1})\|_*$$
$$= \sup_{q \in S_t} \|\delta_q - \delta'_q\|_* \leq \sup_{q \in S_t} \|\delta_q\|_* + \sup_{q \in S_t} \|\delta'_q\|_*$$

We proved that $\|\delta_i\|_* \leq (k+1)i^{k-1}(G+H\|w_i-x_{i-1}\|)$ and $\|\delta'_i\|_* \leq (k+1)i^{k-1}(G+H\|w'_i-x'_{i-1}\|)$ (Lemma 15). Note that $w_i = w'_i$ and $x_i = x'_i$ for all $i \leq t$. Also note that $i \leq t$ for all $i \in S_t$, so:

$$\leq 2(k+1)t^{k-1}(G + H\max_{i\in[t]}\|w_i - x_{i-1}\|).$$

Since $U \leq \log_2(2T)$ and σ_t^2 as defined in (6) satisfies the condition $\sigma_t^2 \geq \Delta_t(\hat{f}_{1:t-1})\log_2(2T)/\rho^2$, Theorem 17 and post-processing imply that $\{g_t + \gamma_t\}_{t=1}^T$ is $(\alpha, \alpha\rho^2/2)$-RDP (so is Algorithm 1). $\quad\square$

C Further Discussions about Differential Privacy

C.1 Example of RDP Distribution

In this subsection, we prove that the multivariate Gaussian distribution $\mathcal{N}(0, I)$ is a (d, α)-RDP distribution w.r.t. 2-norm on $\mathbb{R}^d$ for all $\alpha > 0$ (Definition 4). Namely, $\mathcal{N}(0, I)$ satisfies the following three properties: let $R \sim \mathcal{N}(0, I)$, then (i) $\mathbb{E}[R] = 0$, (ii) $\mathbb{E}[\|R\|_2^2] \leq d$, and (iii) for all $\rho > 0$ and $\mu, \mu' \in \mathbb{R}^d$, if $\sigma^2 \geq \|\mu - \mu'\|_2^2/\rho^2$ then $D_\alpha(P\|Q) \leq \alpha\rho^2/2$, where P, Q denote the distribution of $\sigma R + \mu$ and $\sigma R + \mu'$ respectively.

The first property follows immediately from the definition of $\mathcal{N}(0, I)$. For the second property, $R = (r_1, \ldots, r_d)$ where $r_i \sim N(0, 1)$ iid., so $\mathbb{E}[\|R\|_2^2] = \sum_{i=1}^d \mathbb{E}[r_i^2] = d$. To check the third property, we need the following lemma:

Lemma 18. $D_\alpha(\mathcal{N}(\mu, \sigma^2 I)\|\mathcal{N}(\mu', \sigma^2 I)) = \alpha\|\mu - \mu'\|^2/2\sigma^2$.

Consequently, for all $\sigma^2 \geq \|\mu - \mu'\|_2^2/\rho^2$, $D_\alpha(\mathcal{N}(0, \sigma^2 I)\|\mathcal{N}(\mu, \sigma^2 I)) \leq \alpha\rho^2/2$. This proves that $\mathcal{N}(0, I)$ is indeed a (d, α)-RDP distribution.

Proof of Lemma 18. The density of $\mathcal{N}(\mu, \sigma^2 I)$ is $(2\pi\sigma^2)^{-d/2}\exp(-\|x-\mu\|_2^2/2\sigma^2)$. For short we denote $A = (2\pi\sigma^2)^{-d/2}$ and $B = 1/2\sigma^2$. Then

$$D_\alpha(\mathcal{N}(\mu, \sigma^2 I)\|\mathcal{N}(\mu', \sigma^2 I))$$
$$= \frac{1}{\alpha-1}\log\left(\int_{\mathbb{R}^d} A^\alpha \exp(-B\alpha\|x-\mu\|_2^2)A^{1-\alpha}\exp(-B(1-\alpha)\|x-\mu'\|_2^2)\,dx\right)$$
$$= \frac{1}{\alpha-1}\log\left(\int_{\mathbb{R}^d} A\exp\left(-B(\alpha\|x-\mu\|_2^2 + (1-\alpha)\|x-\mu'\|_2^2)\right)dx\right).$$

Next, observe that

$$x - \mu = (x - \alpha\mu - (1-\alpha)\mu') - (1-\alpha)(\mu - \mu'),$$
$$x - \mu' = (x - \alpha\mu - (1-\alpha)\mu') + \alpha(\mu - \mu').$$

Consequently, upon expanding out $\|x - \mu\|_2^2, \|x - \mu'\|_2^2$, we get:

$$\alpha\|x-\mu\|_2^2 + (1-\alpha)\|x-\mu'\|_2^2 = \|x - \alpha\mu - (1-\alpha)\mu'\|_2^2 + \alpha(1-\alpha)\|\mu - \mu'\|_2^2.$$

Note that $A \exp(-B\|x - \alpha\mu - (1-\alpha)\mu'\|_2^2)$ is the density of $\mathcal{N}(\alpha\mu + (1-\alpha)\mu', \sigma^2 I)$, so it integrates to 1. Therefore,

$$
\begin{aligned}
&D_\alpha(\mathcal{N}(\mu, \sigma^2 I) \| \mathcal{N}(\mu', \sigma^2 I)) \\
&= \frac{1}{\alpha - 1} \log \left(\int_{\mathbb{R}^d} A \exp(-B\|x - \alpha\mu - (1-\alpha)\mu'\|_2^2) \exp(-B\alpha(1-\alpha)\|\mu - \mu'\|_2^2) \, dx \right) \\
&= \frac{1}{\alpha - 1} \log \left(\exp(-B\alpha(1-\alpha)\|\mu - \mu'\|_2^2) \right) = B\alpha\|\mu - \mu'\|_2^2.
\end{aligned}
$$

Recall that $B = 1/2\sigma^2$, and this completes the proof. $\qquad\square$

C.2 Extension to Pure-DP Mechanisms

In the main text, we focus on Renyi differential privacy, and we defined RDP-distribution (Definition 4) accordingly. We can always extend our result in a pure differential privacy setting.

Definition 5 (V-DP distribution). A distribution $\mathcal{D}$ on $\mathbb{R}^d$ is said to be a *DP distribution on norm* $\|\cdot\|$ *with variance constant* V (or simply $\mathcal{D}$ is V-DP on $\|\cdot\|$) if it satisfies that for $R \sim \mathcal{D}$ (i) $\mathbb{E}[R] = 0$, (ii) $\mathbb{E}[\|R\|_*^2] \leq V$, and (iii) for all $\epsilon > 0$ and $\mu, \mu' \in \mathbb{R}^d$, if $\sigma^2 \geq \|\mu - \mu'\|_*^2/\epsilon^2$, then $p((x - \mu)/\sigma)/p((x - \mu')/\sigma) \leq \exp(\epsilon)$ for all $x \in \mathbb{R}^d$, where $p(x)$ is the density of $\mathcal{D}$.

The tree aggregation described in Appendix B also works for pure DP mechanisms as well. Therefore, if we assume $\mathcal{D}$ in Algorithm 1 with a V-DP distribution and change the definition of σ_t^2 in (6) correspondingly, Algorithm 1 can be modified to an purely ϵ-DP mechanism.

Next, we can show that exponential mechanism in general norm satisfies this definition:

Theorem 19. *Consider a probability density* $p(x) = A \exp(-\|x\|_*)$ *on* $(\mathbb{R}^d, \|\cdot\|)$, *where* A *is some normalization constant. Also define* $V = \int_{\mathbb{R}^d} \|x\|_*^2 A \exp(-\|x\|_*) \, dx$. *Then distribution* $\mathcal{D}$ *with density* p *is a* $\mathcal{K}$-DP *distribution.*

Proof. Let $R \sim \mathcal{D}, \mu, \mu' \in \mathbb{R}^d$, and $\sigma^2 \geq 0$. Since the density p is symmetric, $\mathbb{E}[R] = 0$; and by definition, $\mathbb{E}[\|R\|_*^2] = V$. For the third property,

$$
\frac{p((x - \mu)/\sigma)}{p((x - \mu')/\sigma)} = \frac{A \exp(-\|x - \mu\|_*/\sigma)}{A \exp(-\|x - \mu'\|_*/\sigma)} = \exp\left(\frac{-\|x - \mu\|_* + \|x - \mu'\|_*}{\sigma} \right)
$$

By triangular inequality, $-\|x - \mu\|_* + \|x - \mu'\|_* \leq \|\mu - \mu'\|_*$, so:

$$
\leq \exp\left(\frac{\|\mu - \mu'\|_*}{\sigma} \right).
$$

Hence, for all $\sigma \geq \|\mu - \mu'\|_*/\epsilon$, this is further bounded by $\exp(\epsilon)$. $\qquad\square$

D Proofs for the Optimistic Case (Section 3)

Theorem 6. *Suppose Assumption 1 - 4 hold, and* $\mathcal{D}$ *is a* (V, α)-RDP *distribution. Set* $\beta_t = t^k$ *and* σ_t^2 *as defined in* (6), *If the online learner is optimistic (satisfying* (8)) *with* t-th *gradient* $\bar{g}_t = g_t + \gamma_t$ *and* t-th *hint* $\hat{g}_t = \bar{g}_{t-1}$, *then*

$$
\frac{\mathbb{E}[\text{Regret}_T(x^*)]}{\beta_{1:T}} \leq O\left(D(G + DH)\left(1 + \frac{\sqrt{V} \log T}{\sqrt{\lambda}\rho} \right) \frac{1}{T^{3/2}} \right).
$$

Proof. Recall that $g_t = \sum_{i=1}^t \delta_i$, then

$$
\|\bar{g}_t - \hat{g}_t\|_*^2 = \|\delta_t + \gamma_t - \gamma_{t-1}\|_*^2 \leq 3\|\delta_t\|_*^2 + 3\|\gamma_t\|_*^2 + 3\|\gamma_{t-1}\|_*^2.
$$

We showed (Lemma 15) that

$$
\|\delta_t\|_* \leq (k+1)(G + H\|w_t - x_{t-1}\|)t^{k-1} \leq (k+1)(G + DH)t^{k-1}.
$$

Also recall the bound of $\mathbb{E}[\|\gamma_t\|_*^2]$ in (7), so:

$$\mathbb{E}[\|\bar{g}_t - \hat{g}_t\|_*^2] \leq 3(k+1)^2(G+DH)^2 t^{2k-2} + \frac{48(k+1)^2V(G+DH)^2}{\lambda\rho^2}\log_2^2(2T)t^{2k-2}$$

$$= 3(k+1)^2(G+DH)^2 t^{2k-2}\left(1 + \frac{16V\log_2^2(2T)}{\lambda\rho^2}\right).$$

Recall that $\mathbb{E}[\text{Regret}_T(x^*)] \leq O(\mathbb{E}[D\sqrt{\sum_{t=1}^T \|\bar{g}_t - \hat{g}_t\|_*^2}])$. By Jensen's inequality,

$$\mathbb{E}\left[D\sqrt{\sum_{t=1}^T \|\bar{g}_t - \hat{g}_t\|_*^2}\right] \leq D\sqrt{\sum_{t=1}^T \mathbb{E}[\|\bar{g}_t - \hat{g}_t\|_*^2]}$$

$$\leq \sqrt{3}(k+1)D(G+DH)\left(1 + \frac{4\sqrt{V}\log_2(2T)}{\sqrt{\lambda}\rho}\right)T^{k-1/2}.$$

Finally, dividing this bound by $\beta_{1:T} \geq T^{k+1}/(k+1)$ completes the proof. $\qquad\square$

E Proofs for the Strongly Convex Case (Section 4)

Lemma 20. *For any sequence $\beta_t > 0, g_t \in \mathbb{R}^d$, suppose an online learner predicts w_t and receives t-th loss $\ell_t(w) = \langle g_t, w\rangle$, and define $x_t = \sum_{i=1}^t \frac{\beta_i w_i}{\beta_{1:t}}$. If $\mathcal{L}$ is μ-strongly convex w.r.t. $\|\cdot\|$, then*

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \leq \text{Regret}_T(x^*) + \sum_{t=1}^T\left(\langle\beta_t\nabla\mathcal{L}(x_t) - g_t, w_t - x^*\rangle - \frac{\beta_t\mu}{2}\|x_t - x^*\|^2\right).$$

Proof. We start with the strong convexity identity $\mathcal{L}(x^*) \geq \mathcal{L}(x_t) + \langle\nabla\mathcal{L}(x_t), x^* - x_t\rangle + \frac{\mu}{2}\|x_t - x^*\|^2$:

$$\sum_{t=1}^T \beta_t(\mathcal{L}(x_t) - \mathcal{L}(x^*)) \leq \sum_{t=1}^T \beta_t\langle\nabla\mathcal{L}(x_t), x_t - x^*\rangle - \frac{\beta_t\mu}{2}\|x_t - x^*\|^2. \qquad (12)$$

With the same argument in the proof of Lemma 1, we can show:

$$\beta_t\langle\nabla\mathcal{L}(x_t), x_t - x^*\rangle = \beta_t\langle\nabla\mathcal{L}(x_t), x_t - w_t\rangle + \beta_t\langle\nabla\mathcal{L}(x_t), w_t - x^*\rangle$$

Recall the definition that $\beta_{1:t}x_t = \beta_{1:t-1}x_{t-1} + \beta_t w_t$ and thus $\beta_t(x_t - w_t) = \beta_{1:t-1}(x_{t-1} - x_t)$. Also, since $\mathcal{L}$ is convex, $\langle\nabla\mathcal{L}(x_t), x_{t-1} - x_t\rangle \leq \mathcal{L}(x_{t-1}) - \mathcal{L}(x_t)$, so:

$$\leq \beta_{1:t-1}\mathcal{L}(x_{t-1}) - \beta_{1:t-1}\mathcal{L}(x_t) + \langle\beta_t\nabla\mathcal{L}(x_t) - g_t + g_t, w_t - x^*\rangle.$$

Consequently, moving $\sum_{t=1}^T \beta_t\mathcal{L}(x_t)$ to the right side and taking the telescopic sum in (12) gives:

$$-\beta_{1:T}\mathcal{L}(x^*) \leq \sum_{t=1}^T \beta_t\langle\nabla\mathcal{L}(x_t), x_t - x^*\rangle - \beta_t\mathcal{L}(x_t) - \frac{\beta_t\mu}{2}\|x_t - x^*\|^2$$

$$\leq -\beta_{1:T}\mathcal{L}(x_T) + \text{Regret}_T(x^*) + \sum_{t=1}^T\langle\beta_t\nabla\mathcal{L}(x_t) - g_t, w_t - x^*\rangle - \frac{\beta_t\mu}{2}\|x_t - x^*\|^2.$$

Moving $\beta_{1:T}\mathcal{L}(x_T)$ to the left completes the proof. $\qquad\square$

This lemma immediately implies Lemma 7.

Lemma 7. *Suppose $\mathcal{L}$ is μ-strongly convex w.r.t. $\|\cdot\|$. If we replace $\ell_t(w) = \langle g_t + \gamma_t, w\rangle$ with $\bar{\ell}_t(w) = \ell_t(w) + \frac{\beta_t\mu}{4}\|w - x_t\|^2$ in Algorithm 1, and denote the associated regret by $\overline{\text{Regret}}_T$, then*

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^T\langle\beta_t\nabla\mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^*\rangle - \frac{\beta_t\mu}{8}\|w_t - x^*\|^2$$

$$\leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^T \frac{2\|\beta_t\nabla\mathcal{L}(x_t) - g_t - \gamma_t\|_*^2}{\beta_t\mu}.$$

Proof. By definition, $\bar{\ell}_t(w) = \ell_t(w) + \frac{\beta_t \mu}{4}\|w - x_t\|^2$, so

$$\overline{\text{Regret}}_T(x^*) = \sum_{t=1}^T \left(\ell_t(w_t) + \frac{\beta_t \mu}{4}\|w_t - x_t\|^2 \right) - \left(\ell_t(x^*) + \frac{\beta_t \mu}{4}\|x^* - x_t\|^2 \right)$$

$$= \text{Regret}_T(x^*) + \sum_{t=1}^T \frac{\beta_t \mu}{4}(\|w_t - x_t\|^2 - \|x_t - x^*\|^2).$$

Upon substituting this equation into Lemma 20, we get:

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*))$$

$$\leq \overline{\text{Regret}}_T(x^*) - \sum_{t=1}^T \frac{\beta_t \mu}{4}(\|w_t - x_t\|^2 - \|x_t - x^*\|^2)$$

$$+ \sum_{t=1}^T \langle \beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^* \rangle - \frac{\beta_t \mu}{2}\|x_t - x^*\|^2$$

$$\leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^T \langle \beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t, w_t - x_t \rangle - \frac{\beta_t \mu}{4}(\|w_t - x_t\|^2 + \|x_t - x^*\|^2)$$

$$\leq \overline{\text{Regret}}_T(x^*) + \sum_{t=1}^T \langle \beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^* \rangle - \frac{\beta_t \mu}{8}\|w_t - x^*\|^2.$$

The last inequality follows from the identity $\|w_t - x^*\|^2 \leq 2\|w_t - x_t\|^2 + 2\|x_t - x^*\|^2$.

For the second inequality in the lemma, by Fenchel-Young's inequality,

$$\langle \beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^* \rangle - \frac{\beta_t \mu}{8}\|w_t - x^*\|^2$$

$$\leq \|\beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t\|_* \|w_t - x^*\| - \frac{\beta_t \mu}{8}\|w_t - x^*\|^2$$

For any quadratic of form $ax - bx^2$ and $a, b > 0$, note that $\sup_x ax - bx^2 \leq a^2/4b$, so:

$$\leq \frac{2\|\beta_t \nabla \mathcal{L}(x_t) - g_t - \gamma_t\|_*^2}{\beta_t \mu}.$$

$\square$

Proposition 21. *Suppose W is a convex bounded domain with diameter D, and let $u \in W$ and $f(w) = \|w - u\|^2$. Then for all $w \in W$ and $v \in \partial f(w)$, $\|v\|_* \leq 2D$.*

Proof. Let $\phi(r) = r^2$ and $g(w) = \|w - u\|$, then $f(w) = \phi \circ g(w)$. By chain rule of sub-differentials (Corollary 16.72 Bauschke et al. (2011)),

$$\partial f(w) = \{\alpha v' : \alpha \in \partial \phi(g(w)), v' \in \partial g(w)\}$$

$$= \{2\|w - u\|v' : v' \in \partial\|w - u\|\}.$$

By assumption, $\|w - u\| \leq D$. Moreover, $\|\cdot\|$ is 1-Lipschitz (because $\|x\| - \|y\| \leq \|x - y\|$), so $\|v'\|_* \leq 1$ for all $v' \in \partial\|w - u\|$. As a result, for all $v \in \partial f(w)$, $\|v\|_* = 2\|w - u\|\|v'\|_* \leq 2D$. $\square$

F   Proofs for the Parameter-free Case (Section 5)

Definition 6. A random vector $X \in \mathbb{R}^d$ is said to be σ-norm-sub-Gaussian, denoted by $\text{nSG}(\sigma)$ if

$$\mathcal{P}\{\|X - \mathbb{E}[X]\|_2 \geq \epsilon\} \leq 2\exp\left(-\frac{\epsilon^2}{2\sigma^2}\right), \forall \epsilon.$$

We will rely on the following concentration bound on norm-sub-Gaussian random vectors.

Lemma 22 (Lemma 1, Jin et al. (2019))**.** *There exists a universal C such that (i) if $\|X\| \leq \sigma$, then X is $\mathrm{nSG}(C\sigma)$ nad (ii) if X is σ-sub-Gaussian, then X is $\mathrm{nSG}(C\sqrt{d}\sigma)$.*

Lemma 23 (Corollary 8, Jin et al. (2019))**.** *There exists a universal constant C such that if $X_i|X_{1:i-1}$ is mean-zero $\mathrm{nSG}(\sigma_i)$ for all $X_1, \ldots, X_t$, then for any fixed $\delta > 0$ and $B > b > 0$ such that $b < \sum_{i=1}^t \sigma_i^2 \leq B$ almost surely, with probability at least $1 - \delta$,*

$$\left\| \sum_{i=1}^t X_i \right\|_2 \leq C \sqrt{\sum_{i=1}^t \sigma_i^2 \left(\log \frac{2d}{\delta} + \log\log \frac{B}{b} \right)}.$$

Recall that $\beta_t \nabla \mathcal{L}(x_t) - g_t = \sum_{i=1}^t X_i$ (Lemma 15), where

$$X_i = [\beta_i \nabla \mathcal{L}(x_i) - \beta_{i-1}\nabla\mathcal{L}(x_{i-1})] - [\beta_i \nabla\ell(x_i, z_i) - \beta_{i-1}\nabla\ell(x_{i-1}, z_i)];$$

and $\gamma_t = \sum_{i \in I_t} R_i$, where $R_i = \sigma_i \tilde{R}_i$ and $R_i \sim \mathcal{D}$ i.i.d. We have the following lemma:

Lemma 24. *Suppose Assumption 2 - 4 hold w.r.t. the 2-norm, and suppose $\mathcal{D}$ is a (V, α)-RDP distribution and is $\sigma_{\mathcal{D}}$-sub-Gaussian, i.e.,*

$$\mathcal{P}\{ \sup_{\|a\|=1} \langle X, a \rangle \geq \epsilon \} \leq \exp\left(-\frac{\epsilon^2}{2\sigma_{\mathcal{D}}^2} \right).$$

Also set $\beta_t = t^k$. Then there exists a universal constant C such that $X_i|X_{1:i-1}$ are mean-zero $\mathrm{nSG}(\sigma_{X_i})$ and $R_i|R_{1:i-1}$ are mean-zero $\mathrm{nSG}(\sigma_{R_i})$ for all i, where

$$\sigma_{X_i} = 2C(k+1)(G + H\|w_i - x_{i-1}\|_2)i^{k-1},$$
$$\sigma_{R_i} = C\sqrt{d}\sigma_{\mathcal{D}}\sigma_i.$$

Proof. Since we assume $\mathbb{E}[\nabla\ell(x, z)] = \nabla\mathcal{L}(x)$ for all x, $\mathbb{E}[X_i|X_{1:i-1}] = 0$. Also, since $\mathcal{D}$ is a (V, α)-RDP distribution (Definition 4) and $R_i = \sigma_i\tilde{R}_i$'s are independent, $\mathbb{E}[R_i|R_{1:i-1}] = \mathbb{E}[\tilde{R}_i] = 0$.

For the second part, in Lemma 15 we proved that

$$\|\delta_i\|_2 = \|\beta_i \nabla\ell(x_i, z_i) - \beta_{i-1}\nabla\ell(x_{i-1}, z_i)\|_2 \leq (k+1)(G + H\|w_i - x_{i-1}\|_2)i^{k-1}.$$

The same bound holds for $\beta_i \nabla\mathcal{L}(x_i) - \beta_{i-1}\nabla\mathcal{L}(x_{i-1})$ following the same argument. Therefore,

$$\|X_i\|_2 \leq 2(k+1)(G + H\|w_i - x_{i-1}\|_2)i^{k-1}.$$

Moreover, since we assume $\tilde{R}_i \sim \mathcal{D}$ is $\sigma_{\mathcal{D}}$-sub-Gaussian, $R_i = \sigma_i\tilde{R}_i$ is $\sigma_i\sigma_{\mathcal{D}}$-sub-Gaussian. Hence, by Lemma 22, $X_i|X_{1:i-1}$ and $R_i|R_{1:i-1}$ are norm-sub-Gaussian. $\qquad\square$

Theorem 10. *Suppose w.r.t. 2-norm, W is bounded by D and ℓ is G-Lipschitz and H-smooth. Suppose $\mathcal{D}$ is (V, α)-RDP distribution and is $\sigma_{\mathcal{D}}$-sub-Gaussian. If we set $\beta_t = t^3$ (i.e. $k = 3$) and set σ_t^2 as defined in (6), then with probability at least $1 - \delta$,*

$$\mathcal{L}(x_T) - \mathcal{L}(x^*) \leq \frac{4}{T^4}\left(\mathrm{Regret}_T(x^*) + \sum_{t=1}^T \xi_t(\|w_t\|_2 + \|x^*\|_2) + \nu_t(\|w_t\|_2^2 + \|x^*\|_2^2) \right).$$

where C is a universal constant, $A = 8\sqrt{2}C^2$, $A' = 8\sqrt{d}\sigma_{\mathcal{D}}C^2$, $\kappa = 1 + DH/G$, and

$$\xi_t = AG\Phi t^{5/2} + A'(G + DH)\frac{\Phi\log_2(2T)t^2}{\rho}, \ \nu_t = 28AH\Phi t^{5/2}, \ \Phi = \sqrt{\log\frac{20dT\log(2\kappa T)}{\delta}}.$$

Proof. We start with Eq. (2):

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \leq R_T(x^*) + \sum_{t=1}^T \langle \beta_t \nabla\mathcal{L}(x_t) - g_t - \gamma_t, w_t - x^* \rangle$$

$$\leq R_T(x^*) + \sum_{t=1}^T (\|\beta_t \nabla\mathcal{L}(x_t) - g_t\|_2 + \|\gamma_t\|_2)(\|w_t - x^*\|_2). \quad (13)$$

Step 1. By Lemma 23 and 24, for each t, with probability $1 - \delta/2T$,

$$\|\beta_t \nabla \mathcal{L}(x_t) - g_t\|_2 \leq C \sqrt{\sum_{i=1}^{t} \sigma_{X_i}^2 \left(\log \frac{4dT}{\delta} + \log \log \frac{B}{b} \right)}.$$

Since we choose $\beta_t = t^3$ (i.e., $k = 3$),

$$\sigma_{X_i} = 8Ci^2(G + H\|w_i - x_{i-1}\|_2).$$

Next, we can bound $\sum_{i=1}^{t} \sigma_{X_i}^2$ as follows: for all t,

$$\sum_{i=1}^{t} \sigma_{X_i}^2 \leq B := 64C^2(G + DH)^2 T^5,$$

$$\sum_{i=t}^{t} \sigma_{X_i}^2 \geq b := \sigma_{X_1}^2 = 64C^2(G + H\|w_1\|_2)^2.$$

Recall that $\kappa = 1 + DH/G$, so

$$\frac{B}{b} = \frac{(G + DH)^2 T^5}{(G + H\|w_1\|_2)^2} \leq (\kappa T)^5.$$

Also recall that $\Phi = \sqrt{\log(20dT \log(2\kappa T)/\delta)}$, so

$$\sqrt{\log \frac{4dT}{\delta} + \log \log \frac{B}{b}} \leq \sqrt{\log \frac{20dT \log(\kappa T)}{\delta}} \leq \Phi.$$

Therefore, with probability at least $1 - \delta/2T$,

$$\|\beta_t \nabla \mathcal{L}(x_t) - g_t\|_2 \leq C\Phi \sqrt{\sum_{i=1}^{t} [8Ci^2(G + H\|w_i - x_{i-1}\|_2)]^2}$$

$$\leq 8C^2\Phi \sqrt{\sum_{i=1}^{t} i^4 (G + H\|w_i - x_{i-1}\|_2)^2}. \tag{14}$$

By union bound, with probability at least $1 - \delta/2$,

$$\sum_{t=1}^{T} \|\beta_t \nabla \mathcal{L}(x_t) - g_t\|_2 \|w_t - x^*\|_2$$

$$\leq 8C^2\Phi \sum_{t=1}^{T} \sqrt{\sum_{i=1}^{t} i^4 (G + H\|w_i - x_{i-1}\|_2)^2} \|w_t - x^*\|_2$$

We use the identity $(a + b)^2 \leq 2a^2 + 2b^2$ and $\sqrt{a + b} \leq \sqrt{a} + \sqrt{b}$:

$$\leq 8C^2\Phi \sum_{t=1}^{T} \left(\sqrt{\sum_{i=1}^{t} 2G^2 i^4} + \sqrt{\sum_{i=1}^{t} 2H^2 \|w_i - x_{i-1}\|_2^2 i^4} \right) \|w_t - x^*\|_2 \tag{15}$$

1.1. We bound these two sums separately. For the first sum, recall that $A = 8\sqrt{2}C^2$, so

$$8C^2 \sum_{t=1}^{T} \sqrt{\sum_{i=1}^{t} 2G^2 i^4} \|w_t - x^*\|_2 \leq AG \sum_{t=1}^{T} t^{5/2} (\|w_t\|_2 + \|x^*\|_2).$$

1.2. For the second sum, we apply Young's inequality ($ab \leq \frac{1}{2\lambda}a^2 + \frac{\lambda}{2}b^2$) for each t:

$$8C^2 \sum_{t=1}^{T} \sqrt{\sum_{i=1}^{t} 2H^2 \|w_i - x_{i-1}\|_2^2 i^4} \|w_t - x^*\|_2$$

$$\leq AH \sum_{t=1}^{T} \frac{1}{2\lambda_t} \sum_{i=1}^{t} (\|w_i - x_{i-1}\|_2^2 i^4) + \frac{\lambda_t}{2} \|w_t - x^*\|_2^2$$

We first bound $\|w_i - x_{i-1}\|_2^2 \leq 2\|w_i\|_2^2 + 2\|x_{i-1}\|_2^2$. Recall that $x_0 = 0$ and for $i \geq 2$, $x_{i-1} = \sum_{j=1}^{i-1} \frac{\beta_j}{\beta_{1:i-1}} w_j$, so $\|x_{i-1}\|_2^2 \leq \sum_{j=1}^{i-1} \frac{\beta_j}{\beta_{1:i-1}} \|w_j\|_2^2$ by convexity. Consequently,

$$\leq AH \sum_{t=1}^{T} \left(\frac{1}{\lambda_t} \sum_{i=1}^{t} i^4 \|w_i\|_2^2 + \frac{1}{\lambda_t} \sum_{i=2}^{t} i^4 \sum_{j=1}^{i-1} \frac{\beta_j \|w_j\|_2^2}{\beta_{1:i-1}} + \lambda_t(\|w_t\|_2^2 + \|x^*\|_2^2) \right). \quad (16)$$

We define $\lambda_t = ct^{5/2}$ for some constant c to be determined later, and we apply change of summation on the first two sums:

Lemma 25. *For any sequence* a_i, b_j, c_k,

$$\sum_{i=1}^{N} a_i \sum_{j=1}^{i} b_j = \sum_{i=1}^{N} b_i \sum_{j=i}^{N} a_j, \quad \text{and} \quad \sum_{i=1}^{N} a_i \sum_{j=1}^{i} b_j \sum_{k=1}^{j} c_k = \sum_{i=1}^{N} c_i \sum_{j=i}^{N} a_j \sum_{k=i}^{j} b_k.$$

1.2.1. For the first summation,

$$\sum_{t=1}^{T} \frac{1}{\lambda_t} \sum_{i=1}^{t} i^4 \|w_i\|_2^2 = \sum_{t=1}^{T} \sum_{i=t}^{T} \frac{1}{\lambda_i} t^4 \|w_t\|_2^2$$

For decreasing function f, $\sum_{i=t+1}^{T} f(i) \leq \int_t^T f(x)\,dx$, then:

$$\leq \sum_{t=1}^{T} \left(\frac{1}{ct^{5/2}} + \int_t^{\infty} \frac{1}{cx^{5/2}}\,dx \right) t^4 \|w_t\|_2^2 \leq \sum_{t=1}^{T} \frac{5}{3c} t^{5/2} \|w_t\|_2^2.$$

1.2.2. For the second term, by Proposition 14, $\beta_{1:i-1} \geq (i-1)^4/4$, so

$$\sum_{t=1}^{T} \frac{1}{\lambda_t} \sum_{i=2}^{t} i^4 \sum_{j=1}^{i-1} \frac{\beta_j \|w_j\|_2^2}{\beta_{1:i-1}} \leq \sum_{t=1}^{T} \frac{1}{ct^{5/2}} \sum_{i=2}^{t} \frac{4i^4}{(i-1)^4} \sum_{j=1}^{i} j^3 \|w_j\|_2^2$$

For all $i \geq 2$, we can bound $i/(i-1) \leq 2$. We then apply change of summation, which gives:

$$\leq \sum_{t=1}^{T} \sum_{i=t}^{T} \frac{1}{ci^{5/2}} \sum_{j=t}^{i} 64t^3 \|w_t\|_2^2 \leq \sum_{t=1}^{T} \frac{192}{c} t^{5/2} \|w_t\|_2^2.$$

The last inequality is again derived from the integral bound:

$$\sum_{i=t}^{T} \frac{1}{i^{5/2}} \sum_{j=t}^{i} 1 \leq \sum_{i=t}^{T} \frac{1}{i^{3/2}} \leq \frac{1}{t^{3/2}} + \int_t^{\infty} \frac{1}{x^{3/2}}\,dx \leq \frac{3}{t^{1/2}}.$$

In conclusion, upon substituting **1.2.1.** and **1.2.2.** into (16) and setting $c = 14$, we get:

$$8C^2 \sum_{t=1}^{T} \sqrt{\sum_{i=1}^{t} 2H^2 \|w_i - x_{i-1}\|_2^2 i^4} \|w_t - x^*\|_2$$

$$\leq AH \sum_{t=1}^{T} \left(\frac{5}{3c} t^{5/2} \|w_t\|_2^2 + \frac{192}{c} t^{5/2} \|w_t\|_2^2 + ct^{5/2}(\|w_t\|_2^2 + \|x^*\|_2^2) \right)$$

$$\leq AH \sum_{t=1}^{T} 28t^{5/2}(\|w_t\|_2^2 + \|x^*\|_2^2).$$

Moreover, upon substituting **1.1.** and **1.2.** into (15), we get: with probability at least $1 - \delta/2$,

$$\sum_{t=1}^{T} \|\beta_t \nabla \mathcal{L}(x_t) - g_t\|_2 \|w_t - x^*\|_2$$

$$\leq A\Phi \sum_{t=1}^{T} Gt^{5/2}(\|w_t\|_2 + \|x^*\|_2) + 28Ht^{5/2}(\|w_t\|_2^2 + \|x^*\|_2^2).$$

Step 2: We can bound $\sum_{t=1}^{T} \|\gamma_t\|_2 \|w_t - x^*\|_2$ in a similar way. By Lemma 24 and definition of σ_t in (6), $R_i | R_{1:i-1}$ is mean-zero $\mathrm{nSG}(\sigma_{R_i})$, where

$$\sigma_{R_i} = C\sqrt{d}\sigma_\mathcal{D}\sigma_i = \frac{8\sqrt{d}\sigma_\mathcal{D}C}{\rho}\sqrt{\log_2(2T)}(G + H\max_{j\in[i]}\|w_j - x_{j-1}\|_2)i^2.$$

Next, we can bound $\sum_{i\in I_t}\sigma_{R_i}^2$ as follows: for all t,

$$\sum_{i\in I_t}\sigma_{R_i}^2 \geq \min_{i\in I_t}\sigma_{R_i}^2 \geq b := \frac{64d\sigma_\mathcal{D}^2C^2}{\rho^2}\log_2(2T)G^2.$$

On the other hand, since $|I_t| \leq \log_2(2T)$,

$$\sum_{i\in I_t}\sigma_{R_i}^2 \leq B_t := \frac{64d\sigma_\mathcal{D}^2C^2}{\rho^2}\log_2^2(2T)(G + DH)^2t^4.$$

Hence, $B_t/b \leq \log_2(2T)\kappa^2T^4 \leq (2\kappa T)^5$ (because $\log_2(2T) \leq 2T$ and $\kappa \geq 1$). By definition of Φ,

$$\sqrt{\log\frac{4dT}{\delta} + \log\log\frac{B_t}{b}} \leq \sqrt{\log\frac{20dT\log(2\kappa T)}{\delta}} = \Phi.$$

Recall that $A' = 8\sqrt{d}\sigma_\mathcal{D}C^2$. By Lemma 23, for each t, with probability at least $1 - \delta/2T$,

$$\|\gamma_t\|_2 \leq C\sqrt{\sum_{i\in I_t}\sigma_{R_i}^2\left(\log\frac{4dT}{\delta} + \log\log\frac{B_t}{b}\right)} \leq \frac{A'}{\rho}(G + DH)\Phi\log_2(2T)t^2. \qquad (17)$$

By union bound, with probability at least $1 - \delta/2$,

$$\sum_{t=1}^{T}\|\gamma_t\|_2\|w_t - x^*\|_2 \leq \sum_{t=1}^{T}\frac{A'}{\rho}(G + DH)\Phi\log_2(2T)t^2(\|w_t\|_2 + \|x^*\|_2).$$

In conclusion, we take the union bound on the results from **step 1.** and **step 2.** and substitute it back to the starting point (13). Then with probability at least $1 - \delta$,

$$\beta_{1:T}(\mathcal{L}(x_T) - \mathcal{L}(x^*)) \leq R_T(x^*) + \sum_{t=1}^{T}28AH\Phi t^{5/2}(\|w_t\|_2^2 + \|x^*\|_2^2)$$

$$+ \sum_{t=1}^{T}\left(AG\Phi t^{5/2} + A'(G + DH)\frac{\Phi\log_2(2T)t^2}{\rho}\right)(\|w_t\|_2 + \|x^*\|_2).$$

Define ξ_t, ν_t as in the theorem, and recall that $\beta_{1:T} \geq T^4/4$. This completes the proof. $\qquad\square$

Lemma 25. *For any sequence a_i, b_j, c_k,*

$$\sum_{i=1}^{N}a_i\sum_{j=1}^{i}b_j = \sum_{i=1}^{N}b_i\sum_{j=i}^{N}a_j, \quad and \quad \sum_{i=1}^{N}a_i\sum_{j=1}^{i}b_j\sum_{k=1}^{j}c_k = \sum_{i=1}^{N}c_i\sum_{j=i}^{N}a_j\sum_{k=i}^{j}b_k.$$

Proof. The proof is basically re-pairing the summations:

$$\sum_{i=1}^{N}\sum_{j=1}^{i}a_ib_j = a_1b_1 + a_2(b_1 + b_2) + a_3(b_1 + b_2 + b_3) + \cdots$$

$$= (a_1 + \cdots + a_N)b_1 + (a_2 + \cdots + a_N)b_2 + \cdots\sum_{i=1}^{T}\sum_{j=0}^{T-i}a_{t-j}b_i.$$

For the second part of the theorem, denote $B_j^i = \sum_{k=j}^i b_k$. By first part,

$$
\begin{aligned}
LHS &= \sum_{i=1}^N a_i \sum_{j=1}^i c_j \sum_{k=j}^i b_k \\
&= \sum_{i=1}^N \sum_{j=1}^i a_i c_j (B_j^N - B_{i+1}^N) \\
&= \sum_{i=1}^N \sum_{j=i}^N a_j c_i B_i^N - a_j c_i B_{j+1}^N \\
&= \sum_{i=1}^N \sum_{j=i}^N a_j c_i B_i^j.
\end{aligned}
$$

We then recover the lemma once we substitute $B_i^j = \sum_{k=i}^j b_k$. $\qquad\square$