Does Momentum Help in Stochastic Optimization? A Sample Complexity Analysis.

Abstract

Stochastic Heavy Ball (SHB) and Nesterov's Accelerated Stochastic Gradient (ASG) are popular momentum methods in optimization. While the benefits of these acceleration ideas in deterministic settings are well understood, their advantages in stochastic optimization are unclear. Several works have recently claimed that SHB and ASG always help in stochastic optimization. Our work shows that i.) these claims are either flawed or one-sided (e.g., consider only the bias term but not the variance), and ii.) when both these terms are accounted for, SHB and ASG do not always help. Specifically, for *any* quadratic optimization, we obtain a lower bound on the sample complexity of SHB and ASG, accounting for both bias and variance, and show that the vanilla SGD can achieve the same bound.

1 INTRODUCTION

In deterministic convex optimization (when one has access to exact gradients), Gradient Descent (GD) is a popular optimization algorithm Cauchy [1847]. In practice, though, exact gradients are not available and one has to rely on their noisy estimates. This brings forth the idea of Stochastic Gradient Descent (SGD). For GD, two classic momentum ideas are used to speed up its convergence: Heavy Ball (HB) [Polyak, 1964, 1987, Qian, 1999] and Nesterov's Accelerated Gradient (NAG) [Nesterov, 1983, 2014, 2005]. Naturally, these momentum-based methods and their variants have also gained significant interest in stochastic settings [Sutskever et al., 2013, Nitanda, 2014a, Hu et al., 2009a]. However, our work shows that the stochastic variants of HB and NAG, i.e., the Stochastic Heavy Ball (SHB) and Nesterov's Accelerated Stochastic Gradient (ASG), are not better than the vanilla SGD for *any* quadratic optimization.

Specifically, we show that the sample complexities[1] of SHB, ASG, and SGD are all of the same order.

We elaborate on the discussion above. The benefit of using momentum in (deterministic) quadratic optimization is the following. Suppose the driving matrix has condition number κ. Then, for any $\epsilon > 0$, GD with an optimal constant stepsize [2] converges to an ϵ-close solution in $\mathcal{O}(\kappa \log \frac{1}{\epsilon})$ iterations. In contrast, both HB and NAG with optimal stepsize and momentum parameters only need $\mathcal{O}(\sqrt{\kappa} \log \frac{1}{\epsilon})$ steps; see, e.g., [Recht, 2010].

Our main claim here is that momentum will not lead to similar advantages in stochastic settings. We use Figure 1 to provide an intuitive justification for this claim. The setup is as follows. We consider a quadratic optimization problem (see Section E for the details) and ensure that only a noisy estimate of its gradient is available in each iteration. This problem is solved using SGD, SHB, and ASG and the three panels show how the Mean Squared Error (MSE) decays for different stepsize and momentum parameter choices. Note that these parameters, once chosen, are fixed, i.e., they do not change from one iteration to the other.

In stochastic settings, the MSE error at any time instance for each of SGD, SHB, and ASG can be broken down into two components: bias and variance. The bias dictates how fast the distance of the initial estimate to the solution is forgotten, while the variance represents a cumulative effect of the noise seen so far. When constant stepsize and momentum parameters are used, the bias decays exponentially fast while the variance converges to some positive constant; this implies the MSE also converges to this (non-zero) positive constant. Both the rate at which the bias decays and the constant to which the variance converges to are influenced by the stepsize and momentum parameter choices.

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[1] Sample complexity refers to the number of iterations required to reach an ϵ-ball around the solution.

[2] Throughout, we only consider algorithms with constant stepsizes, which are widely popular in practice.

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(a) Stepsize - SGD, SHB, ASG: 0.0025   (b) Stepsize - SGD: 0.02, SHB, ASG: 0.0025   (c) Stepsize - SGD, SHB, ASG: 0.02

Figure 1: Comparison of SGD, SHB, and ASG's performances for a 2D quadratic optimization problem (see Section E for details) for the different stepsize choices given above and ϵ-threshold = 2 (denoted by the black horizontal line).

With the above picture in mind, Figure 1 illustrates our claim that whatever SHB and ASG can achieve, it can be matched by SGD. Figure 1a corresponds to the case where a same stepsize is used in all the three algorithms. In this case, the MSE for the momentum based methods (SHB, ASG) decreases faster initially, but settles at a higher limiting value eventually. Accordingly, one may conjecture that SHB and ASG will have a better sample complexity if the ϵ-threshold for the MSE is set above this limit (one such choice of ϵ in this example is 2). However, Figure 1b shows that SGD enjoys a similar performance for a larger stepsize choice. This time one may conjecture that SHB and ASG's performance can be improved if their stepsizes are also increased similarly. Figure 1c discusses this case when the stepsize for momentum methods is increased to match the new stepsize for SGD. Unfortunately, while MSE for momentum based does decrease faster initially, it also settles at a value that is higher than the threshold that we had set before, i.e., 2.

Related Works: Some recent results [Loizou and Richtárik, 2020, Mou et al., 2020, Assran and Rabbat, 2020, Can et al., 2019] claim that SHB and ASG methods are better than SGD in quadratic or least-squares settings. However, Loizou and Richtárik [2020] needs a strong assumption on noise, which Kidambi et al. [2018, Section 6] claims is information-theoretically impossible even in the simple least squares regression problem. The other results either are based on a one-sided analysis (Can et al. [2019] only consider bias, while ignoring variance) or have a major flaw [Mou et al., 2020, Assran and Rabbat, 2020]; see Appendix A.1—A.3 for details.

On the other hand, there are also a few recent negative results on these momentum methods. For general convex optimization, Yuan et al. [2016] show that SHB and ASG are equivalent to SGD with a rescaled step-size. However, this result requires that the stepsize be sufficiently small and the momentum parameter be away from 1. In Kidambi et al. [2018], Liu and Belkin [2020], for one specific instance of the least squares regression with *vanishing noise*, it is shown that the performance of SHB and ASG cannot be better than

that of SGD. Finally, Zhang et al. [2019] consider SHB for quadratic objectives in the noisy setting as our work and provides upper bounds on the rate at which the objective function decreases. They also argue that rescaled SGD performs as well as SHB and demonstrate it empirically but fall short of rigorously coming up with a lower bound that supports their claim.

SHB and ASG have also been studied in the decreasing step size setting. Ghadimi et al. [2014] had given the first global convergence of SHB for quadratic objectives while Yang et al. [2016], Promsinchai et al. [2020], Orvieto et al. [2019] gave a.s. convergence rates for convex objectives. In [Sebbouh et al., 2021], improved bounds on both SGD and SHB have been provided, as compared to previously known bounds. Hu et al. [2009b], Ghadimi and Lan [2012], Xiao [2009] study Nesterov's momentum under a decreasing stepsize setting and show that though the momentum scheme accelerates the convergence of the iterates in the initial part, the acceleration is lost in the asymptotic regime. Vaswani et al. [2019] study ASG with a decreasing momentum parameter and show a linear convergence to the optimal point. However, the noise at any stationary point vanishes to zero in their setting. Separately, other momentum methods have been studied in Allen-Zhu [2018], Nitanda [2014b], Defazio et al. [2014], Johnson and Zhang [2013], Roux et al. [2012] that achieve a better convergence rate than SGD.

The current literature can thus be summarized as follows.

Research Gap: Existing works on SHB and ASG fall into two groups: i.) positive - where the results claim advantages of these methods over SGD and ii.) negative - where the results claim the opposite. Results in the positive group either have a one-sided or a flawed analysis, while the ones in the negative show that these momentum methods do not help in restricted settings.

Key Contribution: Our work belongs to the negative group: SHB and ASG do not always have an advantage over SGD. Specifically, for *all quadratic optimization problems* with persistent noise (noise variance is sufficiently bounded away

from zero) and any small $\epsilon > 0$, we show that number of iterations needed by SHB and ASG to find an ϵ-optimal solution are of the same order as that of SGD. More technically, we obtain a lower bound on sample complexities of SHB and ASG (Theorem 2.5) and show that these are of the same order as the corresponding upper bound for SGD (Theorem 2.8). Our proof techniques are also significantly different from those used in existing lower bounds such as [Kidambi et al., 2018, Liu and Belkin, 2020]. This is because, under non-vanishing noise, the expected error contains an additional term that cannot be accounted for from their analyses (see Remark 2.7).

2 MAIN RESULTS

We state our main results here that provide lower and upper bounds on the sample complexities of SHB and ASG. We use these bounds along with those of SGD to show that all these methods need a similar effort to find an ϵ-optimal solution. We emphasize that our results apply to the whole family of quadratic optimization.

To provide a unified analysis, we look at a generic update rule that includes as special cases SHB, ASG, and SGD. In particular, we consider the linear stochastic approximation with momentum (LSA-M) iterate given by

$$\begin{aligned}
x_{n+1} &= x_n + \alpha(b - Ax_n + M_{n+1}) \\
&\quad + \eta(I_d - \alpha\beta A)(x_n - x_{n-1}). \quad (1) \\
&= x_n + \alpha(b - A(x_n + \eta\beta(x_n - x_{n-1})) + M_{n+1}) \\
&\quad + \eta(x_n - x_{n-1}), \quad (2)
\end{aligned}$$

where $b \in \mathbb{R}^d, A \in \mathbb{R}^{d \times d}$, and $M_{n+1} \in \mathbb{R}^d$ is noise. Also, I_d refers to the $d \times d$ identity matrix.

Assumption 2.1. *(Driving matrix property) A is symmetric and all its eigenvalues are positive.*

We also denote the the eigenvalues of A by $\lambda_{max} = \lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_d = \lambda_{min}$. Note that LSA-M is equivalent to an SGD (if $\eta = 0$ in (1)). It is also equivalent to SHB (if $\beta = 0$ in (1)), and ASG (if $\beta = 1$ in (2)) with $x_0 = x_{-1}$.

Also, under the above assumption, one would expect the iterates in (1) to go to a neighborhood of $x^* := A^{-1}b$.

We next state two assumptions for (M_n): the first is used in Theorem 2.5, while the other is used in Theorem 2.8 and Corollary 2.9.

Assumption 2.2. *(Noise attributes for Theorem 2.5) (M_n) is a martingale difference sequence with respect to the filtration $(\mathcal{F}_n)$, where $\mathcal{F}_n = \sigma(x_m, M_m; m \leq n)$. Further, $\exists K > 0$ such that $\mathbb{E}[M_{n+1}M_{n+1}^T | \mathcal{F}_n] \succeq KI_d$ a.s. for all $\forall n \geq 0$.*

The notation $A \succeq B$ for $A, B \in \mathbb{R}^d$ is used above to imply that $A - B$ is positive semi-definite.

Assumption 2.3. *(Noise attributes for Theorem 2.8) (M_n) is a martingale difference sequence with respect to the filtration $(\mathcal{F}_n)$, where $\mathcal{F}_n = \sigma(x_m, M_m; m \leq n)$. Further, $\exists K \geq 0$ such that $\mathbb{E}[\|M_{n+1}\|^2 | \mathcal{F}_n] \leq K(1 + \|x_n - x^*\|^2)$ a.s. for all $n \geq 0$.*

Assumptions 2.2 and 2.3 are standard [Mandt et al., 2017, Jastrzębski et al., 2018, Cheng et al., 2020, Borkar, 2008]. The first of these holds if and only if all the eigenvalues of $\mathbb{E}[M_{n+1}M_{n+1}^T | \mathcal{F}_n]$ are bounded from below by K, i.e., noise is persistent (or non-vanishing) in all directions. On the other hand, Assumption 2.3 requires that the trace of $\mathbf{E}[M_{n+1}M_{n+1}^T | \mathcal{F}_n]$ be bounded from above. This bound can scale with $\|x_n - x^*\|$ and need not vanish near x^*.

Next, we define sample complexity to quantify the effort required by LSA-M to obtain an ϵ-close solution to x^*.

Definition 2.4. *(Sample Complexity). The sample complexity of (1) is the minimum number of iterations n_0 such that the expected error $\mathbb{E}[\|x_n - x^*\|^2] \leq \epsilon \, \forall n \geq n_0$.*

To enable easy comparison between different algorithms, we shall look at the order of their sample complexities. Towards that, we shall use the notation $n_0 \in \Theta(t)$ to imply that there exist constants c_1 and c_2 (independent of t) such that $c_1 t \leq n_0 \leq c_2 t$. The notation $\tilde{\Theta}(t)$ has a similar meaning but hides the dependence on logarithmic terms.

Theorem 2.5. *(**Lower bound on sample complexity**). Consider the LSA-M iterate in (1) and suppose Assumptions 2.1 and 2.2 hold. Then, there exists an $\epsilon' > 0$ such that for any $\epsilon \in (0, \epsilon')$ and for any choice of $\alpha > 0$, $\beta \in [0, 1]$, and $\eta \in [0, 1]$, the expected error $\mathbb{E}[\|x_{n_0} - x^*\|^2] \geq \epsilon$ for $n_0 \in \tilde{\Theta}\left(\frac{K}{\epsilon\lambda_{min}^2}\right)$. Thus, n_0 is a lower bound on the sample complexity of LSA-M.*

Remark 2.6. *As stated below (2), LSA-M includes SHB and Nesterov's ASG method as special cases and, hence, the above result directly applies to them. In fact, this is the first lower bound on SHB and ASG's sample complexities in quadratic optimization.*

Remark 2.7. *The lower bounds in Kidambi et al. [2018] and Liu and Belkin [2020] are obtained by viewing the expected error in SHB and ASG iterates for least squares as update rules of the form $z_{n+1} = Pz_n$ for some matrix P [Kidambi et al., 2018, Appendix A, p 16] and [Liu and Belkin, 2020, Appendix C, p 12]). In particular, they obtain bounds on the eigenvalues of P to get the desired claim. In contrast, the error relations for SHB and ASG methods in our setup (quadratic optimization with persistent noise) have the form $z_{n+1} = Pz_n + \alpha W_n$ for some matrix P and vector W_n (cf. 3). This forces us to develop a new proof technique that jointly looks at both these terms and show that at least one of them remains larger than ϵ for the choice of n_0 given in Theorem 2.5.*

Method	β	η	α
SGD	-	0	$\min\left(\frac{\lambda_{min}}{\frac{3}{4}\lambda_{min}^2 + C^2K}, \frac{\epsilon\lambda_{min}}{4C^2K}, \frac{2}{\lambda_{max}+\lambda_{min}}\right)$
SHB	0	$\left(1 - \frac{\sqrt{\alpha\lambda_{min}}}{2}\right)^2$	$\min\left(\left(\frac{\lambda_{min}^{3/2}}{\frac{3}{8}\lambda_{min}^2+25C^2K}\right)^2, \left(\frac{\epsilon(\lambda_{min})^{3/2}}{200C^2K}\right)^2, \left(\frac{2}{\sqrt{\lambda_{min}}+\sqrt{\lambda_{max}}}\right)^2\right)$
ASG	1	$\frac{\left(1 - \frac{\sqrt{\alpha\lambda_{min}}}{2}\right)^2}{(1-\alpha\lambda_{min})}$	$\min\left(\left(\frac{\lambda_{min}^{3/2}}{\frac{3}{8}\lambda_{min}^2+25C^2K}\right)^2, \left(\frac{\epsilon(\lambda_{min})^{3/2}}{200C^2K}\right)^2, \frac{1}{\lambda_{max}}\right)$

Table 1: Parameter choices for Theorem 2.8. Here $C = 1$ when the matrix A is symmetric and $C = \frac{\sqrt{d}}{\sigma_{\min}(S)\sigma_{\min}(S^{-1})}$ when A is not symmetric, where $\sigma_{min}(\cdot)$ denotes the smallest singular value and S is the matrix that diagonalizes A, i.e., $S^{-1}AS = D$, a diagonal matrix. When A is symmetric, indeed the three parameter choices correspond to SGD, SHB and ASG. However, with a slight abuse of notation, we stick to the same naming convention even when the driving matrix A is not symmetric.

We next state our upper bound on the sample complexity in Theorem 2.8 and Corollary 2.9. Similar bounds already exist in literature when A is assumed to be symmetric and the noise is assumed to be iid with bounded variance (Can et al. [2019], Zhang et al. [2019]). Here, we show that a similar upper bound holds under more general settings: i.) A is not symmetric but is diagonalizable and has real positive eigenvalues, and ii.) the noise is a martingale difference sequence satisfying Assumption 2.3.

Proposition 2.8. *Consider the LSA-M iterate and let Assumptions 2.1 and 2.3 hold. Then, $\forall\epsilon > 0$, there exists a choice of α, β and η in LSA-M (see Table 1 for exact values) such that the expected error $\mathbb{E}[\|x_n - x^*\|^2] \le \epsilon$, $\forall n > n_0$, where,*

(i) $n_0 \in \tilde{\Theta}(\frac{1}{\alpha\lambda_{min}})$, when $\eta = 0$, and

(ii) $n_0 \in \tilde{\Theta}(\frac{1}{\sqrt{\alpha\lambda_{min}}})$, when $\eta > 0$.

From Table 1, we see that for each case α is a minimum of three terms. The first term arises due to the unbounded noise (Assumption 2.3), the second due to the target neighborhood ϵ and the third from the optimal choice of step-size in the deterministic (no noise scenario) case. Since the bound on n_0 provided in Theorem 2.8 is in terms of α, the minimum of the three terms dictates the sample complexity. Note that ϵ only influences the middle term in all the choices of α given in Table 1. Let $\bar{\epsilon} > 0$ be such that, for any $\epsilon \in (0, \bar{\epsilon})$, the value of α equals the middle term in each of the three cases. Then, the following result is immediate.

Corollary 2.9 (Upper bound on sample complexity). *Consider the LSA-M iterate and suppose Assumptions 2.1 and 2.3 hold. Then, for choice of parameters in Table 1, and any $\epsilon \in (0, \bar{\epsilon})$, $\exists n_0 \in \tilde{\Theta}\left(\frac{K}{\epsilon\lambda_{min}^2}\right)$ such that $\mathbb{E}[\|x_n - x^*\|^2] \le \epsilon$,* $\forall n \ge n_0$.

Remark 2.10. *From Corollary 2.9, we see that for small enough $\epsilon > 0$ the upper bounds on the sample complexities of SGD, SHB, and ASG match the lower bound given in Theorem 2.5. In particular, since an upper bound on the sample complexity of SGD matches a lower bound for SHB and ASG, these latter methods do not always outperform SGD from a sample complexity perspective.*

Remark 2.11. *Consider ϵ small enough such that the minimum in choice of α is achieved by the second term in Table 1. For SGD, the step-size $\alpha = \Theta(\frac{\epsilon\lambda_{min}}{K})$ is larger than the choice of step-size for SHB and ASG, $\alpha = \Theta(\frac{\epsilon^2\lambda_{min}^3}{K^2})$. Observe that SGD chooses a larger step-size than SHB and ASG to reach the ϵ ball. Therefore, although momentum methods appear to have a better performance than SGD if the same step-size is chosen, SGD can match this performance by re-scaling its step-size (See Figure 1).*

Remark 2.12. *When the noise is assumed to be bounded, the first term in the choice of α in Table 1 vanishes for all three parameter choices. Under such an assumption, if ϵ is large enough or K is small enough such that the third term in the choice of α is the minimum, then the sample complexity of both SHB and ASG is better than SGD. We emphasize that such improvements are lost when the noise variance is large or the neighbourhood under consideration is small.*

3 PROOF OF THE LOWER BOUND (THEOREM 2.5)

We begin by defining the transformed iterates $\tilde{x}_n = x_n - x^*$ and rewriting (1) as:

$$\tilde{X}_n = P\tilde{X}_{n-1} + \alpha W_n, \qquad (3)$$

where, $\tilde{X}_n \triangleq \begin{bmatrix} \tilde{x}_n \\ \tilde{x}_{n-1} \end{bmatrix}$, $W_n \triangleq \begin{bmatrix} M_n \\ 0 \end{bmatrix}$ and

$$P \triangleq \begin{bmatrix} I_d - \alpha A + \eta(I_d - \alpha\beta A) & -\eta(I_d - \alpha\beta A) \\ I_d & 0 \end{bmatrix}.$$

The bound on Theorem 2.5 will follow by lower bounding the error expression $\mathbb{E}[\|\tilde{X}_n\|^2]$.

The proof can be summarized by the following key steps:

1. Decompose the original 2d-dimensional update in (3) into d two-dimensional updates using a block diagonalization argument.

2. For each of the two-dimensional update rules, obtain a lower bound on the error expression. We do this in three steps:

   (a) Decompose the error into two components: one that captures the impact of the initialization (*bias*), and the other that concerns the effect of the cumulative noise (*variance*).

   (b) Using this decomposition, we derive the desired lower bound on the error for the special case of SHB (i.e., $\beta = 0$ in (3)). The idea is to show that either the *bias* or the *variance* $> \epsilon$.

   (c) Generalize the result to $\beta \in [0, 1]$ case by showing that it can be reduced to the former case.

3. Use these individual lower bounds to obtain a lower bound on the error for the original 2d-dimensional problem.

Next we describe the technical results involved in each of the above steps.

1. **Reducing the $2d$-dimensional updates into d two-dimensional updates:** First observe that when $d = 1$, i.e., $A \equiv \lambda$, $M_n \in \mathbb{R}$, then (3) reduces to :

$$\tilde{X}_n = \begin{pmatrix} 1 + \eta - \alpha\lambda & -\eta' \\ 1 & 0 \end{pmatrix} \tilde{X}_{n-1} + \alpha \begin{pmatrix} M_n \\ 0 \end{pmatrix}$$

where $\eta' = \eta(1 - \alpha\lambda_i\beta)$ (cf. 1). We want to transform the update in (3) to d such blocks. Towards this, we study the transformed iterates:

$$\tilde{Y}_n = Z\tilde{X}_n, \tag{4}$$

$$\text{where } Z = E_{2d \times 2d} \begin{pmatrix} S & 0 \\ 0 & S \end{pmatrix}.$$

Here, S is the matrix that diagonalizes A and $E_{2d \times 2d}$ is a suitably defined permutation matrix.

Let $\tilde{Y}_n = \begin{pmatrix} \tilde{Y}_n^{(1)} \\ \tilde{Y}_n^{(2)} \\ \vdots \\ \tilde{Y}_n^{(d)} \end{pmatrix}$ and $\tilde{M}_n = \begin{pmatrix} \tilde{M}_{n,1} \\ \tilde{M}_{n,2} \\ \vdots \\ \tilde{M}_{n,d} \end{pmatrix} = SM_n$,

where each $\tilde{Y}_n^{(i)} \in \mathbb{R}^2$ and $\tilde{M}_{n,i} \in \mathbb{R}$. With these

definitions, the $\tilde{Y}_n$ updates can be decoupled into d two-dimensional updates as in the following lemma.

Lemma 3.1. *The updates for $\tilde{Y}_n$ in* (4) *decouple into the following d two-dimensional update rules $\forall i \in [d]$:*

$$\tilde{Y}_n^{(i)} = \begin{pmatrix} 1 - \alpha\lambda_i + \eta' & -\eta' \\ 1 & 0 \end{pmatrix} \tilde{Y}_{n-1}^{(i)} + \alpha \begin{pmatrix} \tilde{M}_{n,i} \\ 0 \end{pmatrix}$$

where $\eta' = \eta(1 - \alpha\lambda_i\beta)$

Hereafter, we denote

$$B_i = \begin{pmatrix} 1 - \alpha\lambda_i + \eta' & -\eta' \\ 1 & 0 \end{pmatrix}, \tilde{W}_n^{(i)} = \begin{pmatrix} \tilde{M}_{n,i} \\ 0 \end{pmatrix} \tag{5}$$

To see why Lemma 3.1 holds, observe that from (3) we have,

$$\tilde{Y}_n = Z\tilde{X}_n = ZP\tilde{X}_{n-1} + \alpha ZW_n$$
$$= ZPZ^{-1}\tilde{Y}_{n-1} + \alpha ZW_n = B\tilde{Y}_{n-1} + \alpha\tilde{W}_n,$$

where $B = ZPZ^{-1} \in \mathbb{R}^{2d \times 2d}$ and $\tilde{W}_n = ZW_n \in \mathbb{R}^{2d}$. It can be shown that (see Section 3.1 for details)

$$B = \begin{pmatrix} B_1 & 0 & \cdots & 0 \\ 0 & B_2 & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \cdots & 0 & B_d \end{pmatrix}, \tilde{W}_n = \begin{pmatrix} \tilde{W}_n^{(1)} \\ \tilde{W}_n^{(2)} \\ \vdots \\ \tilde{W}_n^{(d)} \end{pmatrix}$$

which produces the decoupled updates in Lemma(3.1).

2. **Solving for each two dimensional update:** We consider the two dimensional decoupled update given in Lemma 3.1 for a specific i and express the lower bound on the sample complexity with respect to λ_i.

(a) **Decompose the error as a sum of bias and variance:** First observe that the update from Lemma 3.1 can be re-written as:

$$\tilde{Y}_n^{(i)} = B_i^n \tilde{Y}_0^{(i)} + \alpha \sum_{i=0}^{n-1} B_i^{(n-1-i)} \tilde{W}_{i+1}^{(i)}. \tag{6}$$

Taking the norm squared of the above equation gives:

$$\|\tilde{Y}_n^{(i)}\|^2 = \underbrace{\|B_i^n \tilde{Y}_0^{(i)}\|^2}_{I}$$

$$+ \underbrace{2\alpha \left(B_i^n \tilde{Y}_0^{(i)} \right)^T \left(\sum_{j=0}^{n-1} B_i^{(n-1-i)} \tilde{W}_{j+1}^{(i)} \right)}_{II}$$

$$+ \underbrace{\alpha^2 \left(\sum_{j=0}^{n-1} B_i^{(n-1-i)} \tilde{W}_{j+1}^{(i)} \right)^T \left(\sum_{j=0}^{n-1} B_i^{(n-1-i)} \tilde{W}_{j+1}^{(i)} \right)}_{III}$$

$$\tag{7}$$

It can be shown that expectation of term II is 0 and that of term III is $\alpha^2 \mathbb{E}\left[\sum_{j=0}^{n-1} \|B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)}\|^2 \right]$ (See Section 3.2 for details). This leads to the following lemma.

Lemma 3.2. *For the update in* (6) *the error can be decomposed as follows:*

$$\mathbb{E}\|\tilde{Y}_n^{(i)}\|^2 = \underbrace{\|B_i^n \tilde{Y}_0^{(i)}\|^2}_{Bias} + \underbrace{\alpha^2 \mathbb{E}\left[\sum_{j=0}^{n-1} \|B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)}\|^2 \right]}_{Variance}$$

The *bias* and *variance* here correspond to that of the i-th block of the transformed iterates.

(b) Bounding the error for $\beta = 0$:

Using the fact that $\eta' = \eta$ when $\beta = 0$, the update in Lemma 3.1 reduces to:

$$\tilde{Y}_n^{(i)} = \begin{pmatrix} 1 - \alpha\lambda_i + \eta & -\eta \\ 1 & 0 \end{pmatrix} \tilde{Y}_{n-1}^{(i)} + \alpha \begin{pmatrix} \tilde{M}_{n,i} \\ 0 \end{pmatrix},$$

We show that for n_0 iterations as defined in Theorem 2.5, either the *bias* or the *variance* is larger than ϵ. This is established in the following key lemma.

Lemma 3.3. *Let* $\epsilon_i' = \min\left(\frac{K}{32\lambda_i^2}, \frac{(\tilde{x}_0^{(i)})^2}{72} \right)$. *Then for any* $\epsilon \in (0, \epsilon_i')$, *any* $\alpha > 0$, $\beta = 0$, $\eta \in [0,1]$ *and for some* $n_0 \in \tilde{\Theta}\left(\frac{K}{\epsilon \lambda_i^2} \right)$,

either the bias, $\|B_i^{n_0} \tilde{Y}_0^{(i)}\|^2 > \epsilon$

or the variance, $\alpha^2 \mathbb{E}\left[\sum_{j=0}^{n-1} \|B_i^{n-1-j} \tilde{W}_{j+1}^{(i)}\|^2 \right] > \epsilon$.

Lemma 3.3 along with Lemma 3.2 immediately provides a lower bound on the error, i.e., $\mathbb{E}\|\tilde{Y}_{n_0}^{(i)}\|^2 > \epsilon$ for $\beta = 0$. Lemma 3.3 is the core of the lower bound analysis and an explicit proof is provided in Section 3.3.

(c) Extending (b) to the case $\beta \in (0,1]$:

Observe that the general β update rule in Lemma 3.1 is equivalent to the $\beta = 0$ update with η redefined as η'. Now Lemma 3.3 can be re-used if we can ensure that $\eta' \in [0,1]$. This holds when $\alpha\lambda_i \leq 1$. To see this observe that

$$\eta' = \eta(1 - \alpha\lambda_i\beta) > \eta(1 - \beta) \geq 0.$$

The last inequality follows because $\beta, \eta \in [0,1]$. Also, $\eta' \leq 1$ since $\eta, \alpha, \lambda_i, \beta \leq 1$. Invoking Lemmma 3.3 now provides the lower bound on the error.

However if $\alpha\lambda_i > 1$ then $\eta' < 0$ and Lemma 3.3 cannot be invoked. But in this case regardless of the choice of η, using *bias-variance* decomposition from Lemma 3.2 we have,

$$\mathbb{E}\|\tilde{Y}_n^{(i)}\|^2 = \|B_i^n \tilde{Y}_0^{(i)}\|^2 + \alpha^2 \mathbb{E}\left[\sum_{j=0}^{n-1} \|B_i^{n-1-j} \tilde{W}_{j+1}^{(i)}\|^2 \right]$$

$$\overset{(a)}{\geq} \|B_i^n \tilde{Y}_0^{(i)}\|^2 + \alpha^2 K \sum_{j=0}^{n-1} \|B_i^j e_1\|^2$$

$$\overset{(b)}{\geq} \alpha^2 K \overset{(c)}{>} \frac{K}{\lambda_i^2} \overset{(d)}{>} \epsilon.$$

Here (a) follows Lemma 3.4 which essentially gives a further lower bound on the *variance*; (b) follows from non-negativity of norm and lower bounding the sum with the $j = 0$ term and (c) follows from $\alpha\lambda_i > 1$. Finally (d) follows for any $\epsilon < \frac{K}{\lambda_i^2}$ which in turn is smaller than ϵ_i' as defined in Lemma 3.3.

This shows that Lemma 3.3 holds for any $\beta \in [0,1]$ and concludes the proof for the two-dimensional update.

3. Bounding the error for the full iterate: Finally we consider the d-dimensional problem. We invoke Lemma 3.3 for any $\beta \in [0,1]$, $i = d$ and $\lambda_d(A) = \lambda_{min}$ and bound the expected error of the original update for $\tilde{X}_n$:

$$\mathbb{E}[\|\tilde{X}_{n_0}\|^2] = \mathbb{E}[\|Z^{-1}\tilde{Y}_{n_0}\|^2]$$
$$\geq \sigma_{min}^2(Z^{-1})\mathbb{E}[\|\tilde{Y}_{n_0}\|^2]$$
$$= \mathbb{E}[\|\tilde{Y}_{n_0}\|^2] \geq \mathbb{E}[\|\tilde{Y}_{n_0}^{(d)}\|^2] \geq \epsilon,$$

for all $\epsilon \in (0, \epsilon_d')$ and for some

$$n_0 \in \tilde{\Theta}\left(\frac{K}{\epsilon \lambda_{min}^2} \right)$$

Here $\sigma_{min}^2(Z^{-1}) = 1$ because Z is a product of unitary matrices. Also, note that

$$2\max\left(\|\tilde{x}_n\|^2, \|\tilde{x}_{n-1}\|^2 \right) \geq \|\tilde{x}_n\|^2 + \|\tilde{x}_{n-1}\|^2$$
$$= \|\tilde{X}_n\|^2$$

Therefore the lower bound on $\mathbb{E}[\|\tilde{X}_n\|^2]$ is enough to prove Theorem 2.5.

3.1 PROOF OF LEMMA 3.1

We begin by defining the following matrices:

$$D = \begin{pmatrix} \lambda_1 & 0 & 0 & \cdots & 0 \\ 0 & \lambda_2 & 0 & \cdots & 0 \\ \vdots & & & & \\ 0 & 0 & 0 & \cdots & \lambda_d \end{pmatrix},$$

$E_{2d \times 2d}$ is a permutation matrix such that

$$E_{2d \times 2d} \begin{pmatrix} I_d - \alpha D + \eta I_d & -\eta I_d \\ I_{d \times d} & 0_{d \times d} \end{pmatrix} E_{2d \times 2d}^{-1} = B$$

where

$$B = \begin{pmatrix} B_1 & 0 & \dots & 0 \\ 0 & B_2 & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & B_d \end{pmatrix}, B_i = \begin{pmatrix} 1 + \eta - \alpha\lambda_i & -\eta \\ 1 & 0 \end{pmatrix}.$$

We study the transformed iterates $\tilde{Y}_n = Z\tilde{X}_n$, where $Z := E_{2d \times 2d} \begin{pmatrix} S & 0 \\ 0 & S \end{pmatrix}$. Here S is the matrix that diagonalizes A, i.e., $S^{-1}AS = D$. Since $\tilde{X}_n = P\tilde{X}_{n-1} + \alpha W_n$, we get

$$\tilde{Y}_n = Z\tilde{X}_n = ZP\tilde{X}_{n-1} + \alpha ZW_n$$
$$= ZPZ^{-1}\tilde{Y}_{n-1} + \alpha ZW_n = B\tilde{Y}_{n-1} + \alpha ZW_n.$$

Here the last equality follows because

$$ZPZ^{-1} = E_{2d \times 2d} \begin{pmatrix} S & 0 \\ 0 & S \end{pmatrix} P \begin{pmatrix} S^{-1} & 0 \\ 0 & S^{-1} \end{pmatrix} E_{2d \times 2d}^{-1}$$
$$= E_{2d \times 2d} \begin{pmatrix} I_{d \times d} - \alpha D + \eta I_{d \times d} & -\eta I_{d \times d} \\ I_{d \times d} & 0_{d \times d} \end{pmatrix} E_{2d \times 2d}^{-1}$$
$$= B$$

Recall that $\tilde{Y}_n = \begin{pmatrix} \tilde{Y}_n^{(1)} \\ \tilde{Y}_n^{(2)} \\ \vdots \\ \tilde{Y}_n^{(d)} \end{pmatrix}$ and $\tilde{M}_n = \begin{pmatrix} \tilde{M}_{n,1} \\ \tilde{M}_{n,2} \\ \vdots \\ \tilde{M}_{n,d} \end{pmatrix} = SM_n$,

where each $\tilde{Y}_n^{(i)} \in \mathbb{R}^2$ and $\tilde{M}_{n,i} \in \mathbb{R}$. Notice that

$$ZW_n = E_{2d \times 2d} \begin{pmatrix} S & 0 \\ 0 & S \end{pmatrix} \begin{pmatrix} M_n \\ 0 \end{pmatrix}$$
$$= E_{2d \times 2d} \begin{pmatrix} \tilde{M}_n \\ 0 \end{pmatrix} = \begin{pmatrix} \tilde{M}_{n,1} \\ 0 \\ \tilde{M}_{n,2} \\ 0 \\ \vdots \\ \tilde{M}_{n,d} \\ 0 \end{pmatrix}$$

Therefore we have $\forall i \in [d]$,

$$\tilde{Y}_n^{(i)} = B_i \tilde{Y}_{n-1}^{(i)} + \alpha \begin{pmatrix} \tilde{M}_n^{(i)} \\ 0 \end{pmatrix}.$$

3.2 PROOF OF LEMMA 3.2

Recall the error expression in (7) and note that the expectation of term II is 0 since $\mathbb{E}[\tilde{W}_{i+1}^{(i)}] = 0$ from assumption

2.3. Next consider term III:

$$\alpha^2 \left(\sum_{j=0}^{n-1} B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)} \right)^T \left(\sum_{j=0}^{n-1} B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)} \right)$$
$$= \alpha^2 \sum_{j,k} (\tilde{W}_{j+1}^{(i)})^T (B_i^{(n-1-j)})^T B_i^{(n-1-k)} \tilde{W}_{k+1}^{(k)}$$
$$= \underbrace{\alpha^2 \sum_{j,k} (\tilde{W}_{j+1}^{(i)})^T (B_i^{(n-1-j)})^T B_i^{(n-1-k)} \tilde{W}_{k+1}^{(k)}}_{III(a)}$$
$$+ \underbrace{\alpha^2 \sum_j \|B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)}\|^2}_{III(b)}$$

We now show that the expectation of $III(a)$ is 0. Without loss of generality, suppose $j < k$. Then,

$$\mathbb{E}\left[(\tilde{W}_{j+1}^{(i)})^T (B_i^{(n-1-j)})^T B_i^{(n-1-k)} \tilde{W}_{k+1}^{(i)} \right]$$
$$= \mathbb{E}\left[\mathbb{E}\left[(\tilde{W}_{i+1}^{(i)})^T (B_i^{(n-1-i)})^T B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)} | \mathcal{F}_j \right] \right]$$
$$= \mathbb{E}\left[(\tilde{W}_{i+1}^{(i)})^T (B_i^{(n-1-i)})^T B_i^{(n-1-j)} \mathbb{E}[\tilde{W}_{j+1}^{(i)} | \mathcal{F}_j] \right] = 0.$$

Therefore, taking expectation on both sides of (7) gives

$$\mathbb{E}\|\tilde{Y}_n^{(i)}\|^2 = \underbrace{\|B_i^n \tilde{X}_0\|^2}_{I}$$
$$+ \mathbb{E}\Bigg[\underbrace{\alpha^2 \sum_{j=0}^{n-1} \|B_i^{(n-1-j)} \tilde{W}_{j+1}^{(i)}\|^2}_{III(b)} \Bigg] \qquad (8)$$

3.3 PROOF OF LEMMA 3.3

This is the key result in the lower bound proof. Here we outline the main steps involved in proving the result. The detailed proofs of the all auxiliary lemmas are pushed to Appendix B. The exact value of n_0 is specified in (39). Hereafter, we we work with this choice of n_0.

Before we proceed with the main proof, we provide a lower bound on the *variance* term in the following lemma.

Lemma 3.4. *Under assumption 2.3, the variance term can be lower bounded as follows:*

$$\alpha^2 \mathbb{E}\left[\sum_{j=0}^{n_0-1} \|B_i^{n_0-1-j} \tilde{W}_{j+1}^{(i)}\|^2 \right] \geq \alpha^2 K \sum_{j=0}^{n_0-1} \|B_i^j e_1\|^2.$$

Hereafter for convenience we will refer to this lower bound as the variance term. If α and η are such that $\|B_i^{n_0} \tilde{Y}_0^{(i)}\|^2 \geq \epsilon$, then the claim immediately follows for this choice of α and η. We now consider the case where α and η are such that

$\|B_i^{n_0}\tilde{Y}_0^{(i)}\|^2 < \epsilon$. Now we will show that for this choice of α and η, the *variance* will necessarily be greater than ϵ. Let $\mu_+^{(i)}$ and $\mu_-^{(i)}$ be the eigen-values of B_i defined as below:

$$
\begin{aligned}
\mu_+^{(i)} &= \frac{1}{2}\left((1 - \alpha\lambda_i + \eta) + \Delta^{(i)}\right) \\
\mu_-^{(i)} &= \frac{1}{2}\left((1 - \alpha\lambda_i + \eta) - \Delta^{(i)}\right)
\end{aligned}
\tag{9}
$$

where $\Delta^{(i)} = \sqrt{(1 - \alpha\lambda_i + \eta)^2 - 4\eta}$.

The following Lemma provides a lower bound on the *variance* in terms of the eigen values of B_i and the momentum parameter η assuming the *bias* is less than ϵ.

Lemma 3.5. *Let* $\alpha > 0$ *and* $\eta \in [0, 1]$ *such that* $\|B_i^{n_0}\tilde{X}_0\|^2 < \epsilon_i'$. *Then we have the following bound,*

$$
\alpha^2 K \sum_{j=0}^{n_0-1} \|B_i^j e_1\|^2 \geq \frac{\alpha^2 K}{2(1 - \mu_+^2)(1 - \mu_-^2)(1 - \eta)}.
$$

It can be shown that $(1 - \mu_+^2)(1 - \mu_-^2) = \alpha\lambda_i$ in the denominator of the rhs above. Using this we define the following function:

$$
Q(\eta; \alpha, \lambda_i) \equiv \frac{\alpha K}{2\lambda_i(1 - \eta)(1 - \rho(B_i))},
$$

where $\rho(B_i) = |\mu_+^{(i)}|$ is the spectral radius of B_i. Moreover it can also be shown that Q depends on η also via $\rho(B_i)$. Now to obtain a further lower bound on the *variance* we optimize over the choice of η in $Q(\eta; \alpha, \lambda_i)$ and show that

$$
Q(\eta; \alpha, \lambda_i) \geq \frac{K}{16\lambda_i^2}
$$

Combining this with the definition of Q and Lemma 3.5 gives the following bound:

$$
\alpha^2 K \sum_{j=0}^{n_0-1} \left\|B_i^j e_1\right\|^2 \geq \frac{K}{16\lambda^2}(1 - \rho(B_i))
$$

The following lemma proves all the above claims.

Lemma 3.6. *Let* $\alpha > 0$ *and* $\eta \in [0, 1]$ *such that* $\|B_i^{n_0}\tilde{X}_0\|^2 < \epsilon_i'$. *Then we have the following bound* $\alpha^2 K \sum_{j=0}^{n_0-1} \|B_i^j e_1\|^2 \geq \frac{K}{16\lambda_i^2}(1 - \rho(B_i))$.

Lastly, to show that the *variance* is lower bounded by $\epsilon \in (0, \epsilon')$, we need to show that $(1 - \rho(B_i)) \geq \frac{16\lambda_i^2}{K}\epsilon$. The choice of n_0 and the fact that we assumed $\|B_i^{n_0}\tilde{Y}_0^{(i)}\| < \epsilon$ exactly ensures that. The following lemma proves this claim.

Lemma 3.7. *For any* $\epsilon \in (0, \epsilon_i')$, *if* $\|B_i^{n_0}\tilde{Y}_0^{(i)}\| < \epsilon$, *then* $1 - \rho(B_i) \geq \frac{16\lambda_i^2}{K}\epsilon$.

This completes the proof of Lemma 3.3.

4 CONCLUDING REMARKS

In this work, we analyze the sample complexity of SHB and ASG and provide matching lower and upper bounds up to constants and logarithmic terms. More importantly, we show that the same sample complexity bound can be obtained by standard SGD. Our work also calls into question some of the recent positive results in favour of SHB and ASG in the stochastic regime. We show that such improvements do not take into account all the terms involved in the error decomposition, or have major flaws. We emphasize that our results are specifically for SHB and ASG and other variants of momentum methods could offer improvements over SGD [Jain et al., 2018, Liu and Belkin, 2020, Kingma and Ba, 2015, Cutkosky and Orabona, 2019].

Although some other results do question the superiority of SHB and ASG in the stochastic regime, the assumptions and the setting that these works look at do not correspond to those in the positive results. We also emphasize that the negative results either only consider small step-sizes and momentum parameters not close to 1 or provide such results for specific instances in linear regression. In contrast, our work shows that SHB and ASG cannot obtain an improvement in terms of sample complexity (for small neighbourhoods) over SGD for the entire family of quadratic optimization and holds for all step-sizes and all momentum parameters in $[0, 1]$.

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