Collaborative Compressors in Distributed Mean Estimation with Limited Communication Budget

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Abstract

Distributed high dimensional mean estimation is a common aggregation routine used often in distributed optimization methods. Most of these applications call for a communication-constrained setting where vectors, whose mean is to be estimated, have to be compressed before sharing. One could independently encode and decode these to achieve compression, but that overlooks the fact that these vectors are often close to each other. To exploit these similarities, recently Suresh et al., 2022, Jhunjhunwala et al., 2021, Jiang et al, 2023, proposed multiple *correlation-aware compression schemes.* However, in most cases, the correlations have to be known for these schemes to work. Moreover, a theoretical analysis of graceful degradation of these correlation-aware compression schemes with increasing *dissimilarity* is limited to only the ℓ_2-error in the literature. In this paper, we propose four different collaborative compression schemes that agnostically exploit the similarities among vectors in a distributed setting. Our schemes are all simple to implement and computationally efficient, while resulting in big savings in communication. The analysis of our proposed schemes show how the ℓ_2, ℓ_∞ and cosine estimation error varies with the degree of similarity among vectors.

1 Introduction

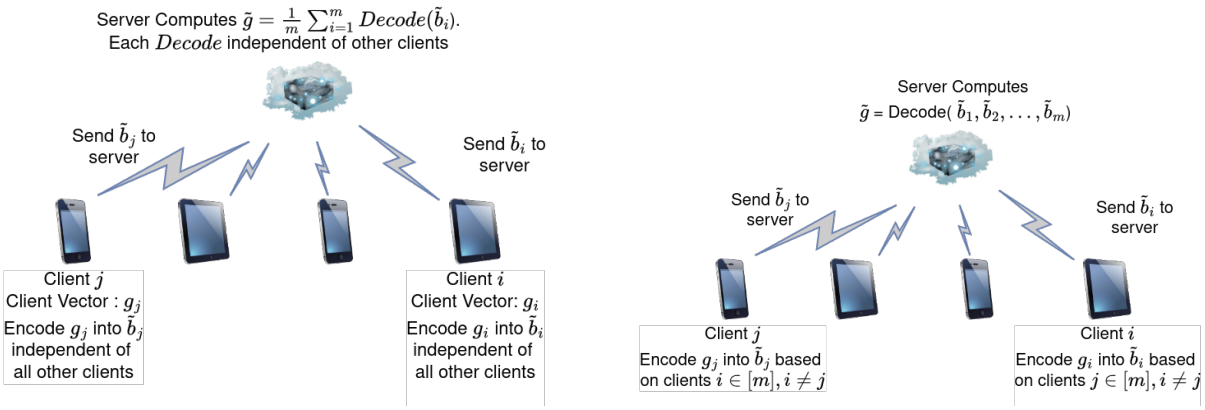
We study the problem of estimating the empirical mean, or average, of a set of high-dimensional vectors in a communication-constrained setup. We assume a distributed problem setting, where m clients, each with a vector $g_i \in \mathbb{R}^d$, are connected to a single server (see, Fig. 1a). Our goal is to estimate their mean g on the server, where

$$g \triangleq \frac{1}{m} \sum_{i \in [m]} g_i. \tag{1}$$

We use $[m]$ to denote the set $\{1, 2, \ldots, m\}$. The clients can communicate with the server via a communication channel which allows limited communication. The server does not have access to data but has relatively more computational power than individual clients.

This problem, referred to as *distributed mean estimation* (DME), is an important subroutine in several distributed learning applications. One common application is distributed training or federated learning (McMahan et al., 2016; McMahan & Ramage, 2017), when different clients correspond to different edge devices.

The typical learning task for DME is supervised learning via gradient-based methods (Bottou & Bousquet, 2007; Robbins & Monro, 1951). In this case, the vectors g_i correspond to the gradient updates for each



(a) Independent Compression

(b) Collaborative Compression

Figure 1: Compression for Distributed Mean Estimation

client i computed on its local training data, and g is the average gradient over all clients. On the other hand, distributed mean estimation is also used in unsupervised learning problems such as distributed KMeans (Liang et al., 2013), distributed PCA (Liang et al., 2014) or distributed power iteration (Li et al., 2021). In distributed KMeans and distributed power iteration, g_i corresponds to estimates of the cluster center and the top eigenvector, respectively, on the ith client.

The naive strategy of clients sending their vectors g_i to the server for DME incurs no error; however, it has a high communication cost, rendering it burdensome in most of the real-world network applications. A principled way to tackle this is to use compression: each client $i \in [m]$ compresses its vector g_i into an efficient encoding $\tilde{b}_i \in \mathcal{B}_i$ which can then be sent to the server; The server forms an estimate $\tilde{g}$ of the mean g using the encodings $\{\tilde{b}_i\}_{i\in[m]}$. We can then compute the error of the estimate $\tilde{g}$ and the number of bits required to communicate $\tilde{b}_i$ (i.e., $\log_2 |\mathcal{B}_i|$) to analyze the efficiency of the compression scheme. As opposed to distributed statistical inference (Braverman et al., 2016; Garg et al., 2014), we do not assume that g_i are sampled from a distribution, and instead the estimation error of these schemes is computed in terms of g_i.

One way to approach this compression paradigm is when each client compresses its vector oblivious to others, and the server separately decodes the vectors before aggregating (Figure 1a). We call this *independent compression* and several existing works (Konečný & Richtárik, 2018; Suresh et al., 2017; Safaryan et al., 2021; Gandikota et al., 2022; Vargaftik et al., 2021; Ghosh et al., 2021) use such a compression scheme. The simplest example of this scheme is RandK (Konečný & Richtárik, 2018), where each client sends only $K \in \mathbb{N}$ coordinates as $\tilde{b}_i$, and the server estimates $\tilde{g}$ as the average of K-sparse vectors from each client. As $K < d$, this scheme requires less communication than sending the full vector g_i from each client $i \in [m]$.

However, independent compressors suffer from a significant drawback, especially when the vectors to be aggregated are similar/not-too-far, which is often the case for gradient aggregation in distributed learning. Consider the case when two distinct clients $i, j \in [m]$ have different vectors $g_i \neq g_j$, but they differ in only one coordinate. Then, independent compressors like RandK will end up sending $\tilde{b}_i$ and $\tilde{b}_j$ which are very similar (in fact, same with high probability) to each other, and therefore wasting communication. Collaborative compressors (Suresh et al., 2022; Szlendak et al., 2021; Jhunjhunwala et al., 2021; Jiang et al., 2023) can alleviate this problem. Figure 1b describes a collaborative compressor, where the encodings $\{\tilde{g}_i\}_{i\in[m]}$ may not be independent of each other and a decoding function *jointly* decodes all encodings to obtain the mean estimate $\tilde{g}$. Clearly, this opens up more possibilities to reduce communication - but also the error of collaborative compressors can be made to scale as the variability of the vectors.

The amount of required communication also depends on the metric for estimation error. Among the existing schemes for collaborative compressors, most provide guarantees on the ℓ_2 error $||\tilde{g} - g||_2^2$ (Suresh et al., 2022; Szlendak et al., 2021; Jhunjhunwala et al., 2021; Jiang et al., 2023). Also, in collaborative compressors, the error must ideally be dependent on *some measure of correlation/distance* among the vectors, which is indeed

Compressor	Error metric	Error	# Bits/client
NoisySign (Algorithm 1)	$\|\|\tilde{g}-g\|\|_\infty$	$\exp(\frac{\|\|g\|\|_\infty^2}{\sigma^2}) \cdot (\Delta_\phi + \sqrt{\frac{\log m}{m}})$	d
HadamardMultiDim (Algorithm 3)	$\mathbb{E}[\|\|\tilde{g}-g\|\|_\infty]$	$\frac{B}{2^{m-1}} + \Delta_{\text{Hadamard}}$	d
SparseReg (Algorithm 4)	$\mathbb{E}[\|\|\tilde{g}-g\|\|_2^2]$	$B^2 \exp\left(-\frac{2m\log L}{d}\right) + \Delta_{\text{reg}}$	$\log L$ ($L \geq 1$ tunable)
OneBit (Algorithm 5)	$\arccos\langle \tilde{g}, g\rangle$	$\pi(\Delta_{\text{corr}} + \frac{d}{mt})$	t ($t \geq 1$ tunable)

Table 1: Theoretical results for our proposed collaborative compression schemes. $\Delta_\Phi, \Delta_{\text{Hadamard}}, \Delta_{\text{reg}}$ and Δ_{corr} are measures of average dissimilarity between vectors $\{g_i\}_{i\in[m]}$ defined in Theorems 1, 2, 3 and Lemma 1 respectively. For NoisySign, $\sigma > 0$ is an algorithm parameter. For HadamardMultiDim, we assume $||g_i||_\infty \leq B, \forall i \in [m]$. For SparseReg, we assume $||g_i||_2 \leq B, \forall i \in [m]$ and L is an algorithm parameter. For OneBit, g is the unit vector along the average $\frac{1}{m}\sum_{i=1}^m g_i$ and $\tilde{g}$ is also a unit vector.

the case for all of these schemes. In this paper, the measure of such a distance is denoted with Δ, with some subscript signifying the exact measure; the vectors in question have high similarity as $\Delta \to 0$. The estimation error naturally grows with the dimension d, and decays with the number of clients m (due to an averaging). One of our major contributions is to design a compression scheme that has significantly improved dependence on the number of clients m to counter the effect of growing dimension d.

If one were to estimate the unit vector in the direction of the average vector $\frac{1}{m}\sum_{i=1}^m g_i$, which is often important for gradient descent applications, using an estimate of the mean with low ℓ_2 error can be highly sub-optimal as the ℓ_2 error might be large even if all the vectors point in the same direction but have different norms. For this, the cosine distance $\arccos(\frac{\langle \tilde{g}, g\rangle}{\|\tilde{g}\|\|g\|})$ is a better measure, which has not been studied in the literature. We also give a compression scheme specifically tailored for this error metric. Another interesting metric is the ℓ_∞-error which has also not been studied except for in (Suresh et al., 2022). There as well, we give an improved dependence of the estimation error on m.

Further drawback of existing collaborative compressors such as, (Jhunjhunwala et al., 2021; Jiang et al., 2023) is that they require the knowledge of correlation between vectors before employing their compression. Without this knowledge, their error guarantees do not hold.

Notation. We use $[a, b]$ to denote $[a, a+1, a+2, \ldots, b]$ for $a, b \in \mathbb{N}$. We use $g^{(j)}, j \in [d]$ to denote the jth coordinate of a vector $g \in \mathbb{R}^d$. For a permutation ρ on $[m]$, $\rho^{(i)}$ denotes mapping of $i \in [m]$ under ρ. We use Π_m to denote the uniform distribution over all permutations of $[m]$. We use $\exp(a)$ to denote $e^{\mathcal{O}(a)}$ for any $a \in \mathbb{R}$. For any matrix $A \in \mathbb{R}^{m\times n}$, A_k for $k \in [m]$ refers to the kth row of A.

Our contributions. We provide four different collaborative compressors, which are communication-efficient, give error guarantees for different error metrics (ℓ_2 error, ℓ_∞ error and cosine distance), and exhibit optimal dependence on the number of clients m and the diameter of ambient space B. To see the advantage of collaboration, we define few natural similarity metrics. All our schemes show graceful degradation of error with the similarity metric between different clients. Our schemes have three subroutines: `Init` which corresponds to initialization/setting up a protocol, `Encode` which is performed individually at each client to obtain their encoding $\tilde{b}_i$, and `Decode` which is performed at the server on all the encodings to obtain estimate of mean $\tilde{g}$.

Below our main contributions are summarized. The theoretical guarantees for our algorithms are provided in Table 1. For each of our proposed compressors, increasing m increases the total communication required (which is a product of m and the # Bits/client), but reduces the estimation error in the appropriate error metric up to a constant Δ that depends on the deviation between client vectors.

1. We provide a simple collaborative scheme based on the popular signSGD (Bernstein et al., 2018a) scheme, NoisySign (Algorithm 1), where sign of each coordinate of a vector is sent after adding Gaussian noise. An advantage of this scheme, compared to others is that we can infer the vector g with an ℓ_∞ error guarantee increasing with $||g||_\infty$ and decreasing with m, without the knowledge of $||g||_\infty$ itself. The dissimilarity is

Compressor	Error	# Bits/client	Notes
RandK (Konečný & Richtárik, 2018)	$\mathcal{O}(\frac{d}{K}\tilde{B}^2)$	$32K + K\log d$	Independent
SRQ (Suresh et al., 2017)	$\mathcal{O}(\frac{\log d}{m(K-1)^2}\tilde{B}^2)$	Kd	Independent
Kashin (Safaryan et al., 2021)	$\mathcal{O}(\left(\frac{10\sqrt{\lambda}}{\sqrt{\lambda}-1)}\right)^4\tilde{B}^2)$	$31 + \lambda d$	Independent
Drive (Vargaftik et al., 2021)	$\mathcal{O}(\tilde{B}^2)$	$32 + d$	Independent
PermK (Szlendak et al., 2021)	$\mathcal{O}((1 - \max\{0, \frac{m-d}{m-1}\})\Delta_2)$	$32K + K\log d$	Collaborative
RandKSpatial (Jhunjhunwala et al., 2021)	$\mathcal{O}(\frac{d}{mK}\Delta_2)$	$32K + K\log d$	Needs Correlation
RandKSpatialProj (Jiang et al., 2023)	$\mathcal{O}(\frac{d}{mK}\Delta_2)$	$32K + K\log d$	Needs Correlation
Correlated SRQ (Suresh et al., 2022)	$\mathcal{O}\left(\frac{1}{m}\min\{\frac{\sqrt{d}\Delta_\infty^d B}{K}, \frac{dB^2}{K^2}\}\right)$	$2d\log K + K\log d$	$\|\|g_i\|\|_2 \le B, \forall i \in [m]$

Table 2: Comparison of existing independent and collaborative compressors in terms of ℓ_2 error and bits communicated. K is the number of coordinates communicated for sparsification methods(RandK, PermK, RandKSpatial, RandKSpatialProj) and the number of quantization levels for quantization methods (SRQ, vqSGD, Correlated SRQ). The constant λ is a parameter of the Kashin scheme. Further, $\tilde{B}^2 = \frac{1}{m}\sum_{i=1}^m ||g_i||_2^2, \Delta_2 = \frac{1}{m}\sum_{i=1}^m ||g_i - g||_2^2$, and $\Delta_\infty = \max_{j\in[d]} \frac{1}{m}\sum_{i=1}^m |g_i^{(j)} - g^{(j)}|$. It is also assumed that a real is equivalent to 32 bits, which is an informal norm in this literature.

$\Delta_\Phi = \mathcal{O}(\frac{1}{m\sigma}\sum_{i=1}^m ||g - g_i||_\infty)$, where σ is the variance of the noise added (Theorem 1). This scheme is described in Section 2.

2. **(ℓ_∞-guarantee)** For vectors with ℓ_∞ norm bounded by B, we propose a collaborative compression scheme, HadamardMultiDim (Algorithm 3) which performs coordinate-wise collaborative binary search. We obtain the best dependence on m and B for the ℓ_∞ error ($\mathcal{O}(B \cdot \exp(-m))$) while suffering from an extra error term Δ_{Hadamard}, which is a measure of average dissimilarity between compressed vectors. Δ_{Hadamard} lies in the range $[\Delta_\infty, \Delta_{\infty,\max}]$ where $\Delta_\infty = \max_{j\in[d]} \frac{1}{m}\sum_{i=1}^m |g_i^{(j)} - g^{(j)}|$ and $\Delta_{\infty,\max} = \max_{j\in[d], i\in[m]} |g_i^{(j)} - g^{(j)}|$ (Theorem 2). In Section 3.3, we provide a practical example where value of Δ_{Hadamard} can be approximated and use it compare theoretical guarantees of HadamardMultiDim with those of baselines in Table 2.

3. **(ℓ_2-guarantee)** For vectors with ℓ_2 norm bounded by B, we provide a collaborative compression scheme SparseReg (Algorithm 4) based on Sparse Regression Codes (Venkataramanan et al., 2014b;a). We obtain the best dependence on B and m for the ℓ_2 error ($\mathcal{O}(B\exp(-m/d))$) while compressing to much less than d bits (in fact, to a constant number of bits) per client. The error consists of a penalty for the dissimilarity, Δ_{reg}, the average dissimilarity between compressed vectors which lies in the range $[\Delta_2, \Delta_{2,\max}]$ where $\Delta_2 = \frac{1}{m}\sum_{i=1}^m ||g - g_i||_2^2$ and $\Delta_{2,\max} = \max_{i\in[m]} ||g - g_i||_2^2$ (see, Theorem 3).

4. **(cosine-guarantee)** For unit norm vectors $\{g_i\}_{i\in[m]}$, we estimate the unit vector g in the direction of the average $\frac{1}{m}\sum_{i=1}^m g_i$. For this, motivated by one-bit compressed sensing (Boufounos & Baraniuk, 2008), our collaborative compression scheme, OneBit (Algorithm 5), sends the sign of the inner product between the vector g_i and a random Gaussian vector. By establishing an equivalence to halfspace learning with malicious noise, we propose two decoding schemes: the first one is based on (Shen, 2023) which is optimal for halfspace learning but harder to implement and a second one, based on (Kalai et al., 2008) which is easy to implement. Both schemes are computationally efficient, and have an extra dissimilarity term in the error, $\Delta_{\text{corr}} = \frac{1}{m\pi}\sum_{i=1}^m \cos^{-1}(\langle g, g_i\rangle)$, which is the appropriate dissimilarity between unit vectors (see Theorem 4).

5. **(Experiments)** We perform a simulation for DME with our schemes as the dissimilarities vary and compare the three different error metrics from above with various existing baselines (Fig 2a-2c). We also used our DME subroutines in the downstream tasks of KMeans, power iteration, and linear regression on real (and federated) datasets (Fig 2d-2i). Our schemes have lowest error in all metrics for low dissimilarity regime.

Organization. In the next subsection, we present related works in distributed mean estimation. The NoisySign algorithm is given in Algorithm 1, and its analysis can be found in Section 2. In Section 3, we present the two schemes obtaining optimal dependence on m, HadamardMultiDim in Subsection 3.1 and

SparseReg in Subsection 3.2. In Section 4, we analyze the OneBit compression scheme. Finally, in Section 5, we provide experimental results for our schemes.

1.1 Related Works

Compressors in Distributed Learning. Starting from (Konečnỳ et al., 2016) most compression schemes in distributed learning involve either quantization or sparsification. In quantization schemes, the real valued input space is quantized to specific levels, and each input is mapped to one of these quantization levels. A theoretical analysis for unbiased quantization was provided in (Alistarh et al., 2017). Subsequently, the distributed mean estimation problem with limited communication was formulated in (Suresh et al., 2017) where two schemes, stochastic rotated quantization (SRQ) and variable length coding, were proposed. These schemes matched the lower bound for communication and ℓ_2 error in terms of $\tilde{B}^2 = \frac{1}{m}\sum_{i=1}^{m}||g_i||_2^2$. Performing a coordinate-wise sign is also a quantization operation, introduced in (Bernstein et al., 2018b). Further advances in quantization include multiple quantization levels (Wen et al., 2017), probabilistic quantization with noise (Chen et al., 2020; Jin et al., 2021; Safaryan & Richtarik, 2021), vector quantization (Gandikota et al., 2022; Zhu et al., 2023), and applying structured rotation before quantization (Vargaftik et al., 2021; Safaryan et al., 2021). Sparsification involves selecting only a subset of coordinates to communicate. Common examples include RandK (Konečný & Richtárik, 2018), TopK (Stich et al., 2018) and their combinations (Beznosikov et al., 2022). Note, for all independent compressors, the ℓ_2 error scales as $\tilde{B}^2$.

Collaborative Compressors. PermK (Szlendak et al., 2021) was the first collaborative compressor, where each client would send a different set of K coordinates. Their error scales with the empirical variance, $\Delta_2 = \frac{1}{m}\sum_{i=1}^{m}||g_i - g||_2^2$. If Δ_2 is known, or one of the vectors g_i is known, the lattice-based quantizer in (Davies et al., 2021) and correlated noise based quantizer in (Mayekar et al., 2021) obtains ℓ_2 error in terms of Δ_2. Further, RandKSpatial (Jhunjhunwala et al., 2021) and RandKSpatialProj (Jiang et al., 2023) utilize the correlation information to obtain the correct normalization coefficients for RandK with rotations, obtaining guarantees in terms of Δ_2. In absence of correlation information, they propose a heuristic. A quantizer also based on correlated noise, was proposed in (Suresh et al., 2022) which achieves the lower bound for scalars. However, for d-dimensional vectors of ℓ_2-norm at most B, their dependence on dimension d and number of clients m can be improved by our schemes.

We provide a summary of existing compressors in Table 2, along with their error guarantees.

2 NoisySign for unbounded $||g_i||_\infty$

In this section, we utilize collaborative compression to use sign compressor when $||g_i||_\infty$ is unbounded. The sign-compressor (Bernstein et al., 2018a) applies the sign function coordinate-wise, where $\text{sign}(x) = +1$ if $x \geq 0$ and -1 otherwise. For this section, we will focus on a single coordinate $j \in [d]$. Note that for any $i \in [m]$, $\text{sign}(g_i^{(j)})$ does not have information about $|g_i^{(j)}|$. Existing compressors (Karimireddy et al., 2020) remedy this by sending $|g_i^{(j)}|$ separately, or assuming that $|g_i^{(j)}|$ is bounded by some constant B (Safaryan & Richtarik, 2021; Jin et al., 2023; Tang et al., 2023). In the second case, the maximum error that can be incurred is $\frac{B}{2}$. This can be improved by adding uniform symmetric noise before taking signs (Chen et al., 2020; Chzhen & Schechtman, 2023) depending on g. However, if no information is available about $|g_i^{(j)}|$, we cannot provide an estimate of $g_i^{(j)}$.

We utilize the concept of adding noise before taking signs, however, to accommodate possibly unbounded $|g_i^{(j)}|$, we add symmetric noise with unbounded support. One choice for such noise is the Gaussian distribution $\mathcal{N}(0, \sigma^2)$ for some $\sigma \in \mathbb{R}$. For $\xi_i^{(j)} \sim \mathcal{N}(0, \sigma^2)$, we send $\tilde{b}_i^{(j)} = \text{sign}(g_i^{(j)} + \xi_i^{(j)})$ as the encoding. Note that $\mathbb{E}[\tilde{b}_i^{(j)}] = \Phi_\sigma(g_i^{(j)})$, where $\Phi_\sigma(t) = 2\Pr_{x\sim\mathcal{N}(0,\sigma^2)}[x \geq -t] - 1 = \text{erf}(\frac{t}{\sqrt{2}\sigma})$, and erf is the error function for the unit normal distribution. A single $\tilde{b}_i^j$ gives us information about $g_i^{(j)}$, however, using it to decode $g_i^{(j)}$ might incur a very large variance. However, assuming that all $g_i^{(j)}$ are close to $g^{(j)}$ for $i \in [m]$, $\frac{1}{m}\sum_{i=1}^{m}\tilde{b}_i^{(j)}$ is a good estimator for $\Phi_\sigma(g^{(j)})$. So, to estimate $g^{(j)}$, we can use $\Phi_\sigma^{-1}(\frac{1}{m}\sum_{i=1}^{m}\tilde{b}_i^{(j)})$. This scheme performed

coordinate-wise is the NoisySign algorithm described in Algorithm 1. We provide estimation error for recovering g using this scheme.

Theorem 1 (Estimation error of noisy sign)**.** *With probability $1 - 2dm^{-c}$, for some constant $c > 0$, the estimation error of Algorithm 1 is*

$$||\tilde{g} - g||_\infty \leq \sqrt{\frac{\pi}{2}} \left(\left(1 - \frac{\Delta_\Phi + \sqrt{\frac{2c \log m}{m}}}{\alpha(||g||_\infty)} \right)^{-1} - 1 \right), \tag{2}$$

where $\Delta_\Phi \triangleq \max_{j \in [d]} |\frac{1}{m} \sum_{i=1}^m \Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})|$ and $\alpha(u) \triangleq 1 - \Phi_\sigma(u)$. Further, for small Δ_ϕ and $\frac{||g||_\infty}{\sigma}$, and large m,

$$||\tilde{g} - g||_\infty = \mathcal{O}\left(\exp(||g||_\infty^2/\sigma^2) \cdot (\Delta_\phi + \sqrt{\log m/m}) \right). \tag{3}$$

Applying Φ_σ^{-1} to estimate g makes our scheme collaborative, as an independent scheme would have individually decoded each client's $\tilde{b}_i^j$ and thus would not take advantage of concentration over m machines. To gain insight into the error, note that $(1-x)^{-1} - 1 \approx x$, for small x. The error increases with the increase in $||g||_\infty$ as we are compressing unbounded variables g_i into the bounded domain $[-1, 1]$ which is the range of the function Φ_σ.The number of clients m determines the resolution with which we can measure on this domain, as the value $\frac{1}{m} \sum_{i=1}^m \tilde{b}_i$ can only be in multiples of $\frac{1}{m}$. Therefore, increasing m decreases the error. As $m \to \infty$, the ℓ_∞-error approaches $\Delta_\Phi \exp(\frac{||g||_\infty}{\sigma^2})$. Note that Δ_Φ determines the average separation between vectors in terms of the Φ_σ operator. If vectors g_i are similar to each other, Δ_Φ is small, and the error is small as a result. Further, Δ_Φ can be bounded by more interpretable quantities if the average separation between g_i and g is small:

$$\Delta_\Phi \leq \sqrt{\frac{2}{\pi}} \frac{1}{m\sigma} \sum_{i \in [m]} ||g_i - g||_\infty. \tag{4}$$

Note that Δ_Φ is always ≤ 2, so if the average error in terms of ℓ_∞ norm is much smaller than σ, then the above bound makes sense. Combining Eq (4) with Eq (3), we can find an asymptotic expression for the estimation error. A large value of σ can help us reduce this upper bound on Δ_ϕ as well as the impact of $\alpha(||g||_\infty)$ on the estimation error. One can tune the value of σ if additional information is known, for instance a large upper bound on $||g||_\infty$ and $\frac{1}{m} \sum_{i=1}^m ||g_i - g||_\infty$. However, even without this information, we can run our algorithm. In contrast, the value of $\max_{i \in [m]} ||g||_\infty$ is required to even run vanilla sign compression algorithms. Additionally, it yields a constant error of $\mathcal{O}(\max_{i \in [m]} ||g_i||_\infty)$, as each sign would need to be accurate. If an upper bound is known for $\max_{i \in [m]} ||g_i||_\infty$, then the error scales with this upper bound. In contrast, for large m and small Δ_Φ, knowledge of a large upper bound helps us to set a large σ and thus our collaborative compressor performs much better. Proofs for this section are provided in Appendix A.

Algorithm 1 NoisySign

`Encode(`g_i`)`
Sample $\xi_i \sim \mathcal{N}(0, \sigma^2 \mathbb{I}_d)$
$\tilde{b}_i = \text{sign}(g_i + \xi_i)$
return $\tilde{b}_i$.
`Decode(`$\{\tilde{b}_i\}_{i \in [m]}$`)`
$\tilde{g}^{(j)} \leftarrow \Phi_\sigma^{-1}(\frac{1}{m} \sum_{i=1}^m \tilde{b}_i^{(j)}), j = 1, \ldots, d$
return $\tilde{g}$

Algorithm 2 Hadamard1DEnc

Input: Scalar s, Level K
$S_K^- = \cup_{k=0}^{2^{K-1}-1} [-B + \frac{2kB}{2^{K-1}}, -B + \frac{(2k+1)B}{2^{K-1}}]$
return -1 if $s \in S_K^-$ else $+1$

Algorithm 3 HadamardMultiDim

`Init()`
Sample a random permutation $\rho \sim \Pi_m$.
Clients and server share ρ.
`Encode(`g_i`)`
for $j \in [d]$ **do**
  $\tilde{b}_i^{(j)} \leftarrow \text{Hadamard1DEnc}(g_i^{(j)}, \rho^{(i)})$
end for
return $\tilde{b}_i$
`Decode(`$\{\tilde{b}_i\}_{i \in [m]}$`)`
for $j \in [d]$ **do**
  $\tilde{g}^{(j)} = \sum_{i=1}^m \tilde{b}_i^{(j)} \cdot \frac{B}{2^{\rho^{(i)}-1}}$
end for
return $\tilde{g}$

For NoisySign, collaboration ensured concentration over m. In the next section, we utilize collaboration to obtain optimal dependence on m.

3 Optimal Dependence on m

If $||g||_\infty$ or $||g||_2$ is bounded, we can obtain an almost optimal exponential decay with m. We provide two schemes that obtain optimal ℓ_∞ (by modifying the sign compressor) and ℓ_2 error dependence in terms of m and the diameter of the space B. The proofs for this section are provided in Appendix B.

3.1 HadamardMultiDim

When the vectors have bounded ℓ_∞ norm, instead of obliviously using the sign compressor on every coordinate on every client, one may be able to divide their range and cleverly select bits to encode the most information. We call our algorithm Hadamard scheme, because the binary-search method involved is akin to the rows of a Hadamard-type matrix.

Assumption 1 (Bounded domain)**.** $||g_i||_\infty \leq B, \forall i \in [m]$.

This would imply that for any $j \in [d]$, $g_i^{(j)} \in [-B, B], \forall i \in [m]$. Now, consider the ith client and the scalar $g_i^{(j)}$ and assume that we are allowed to encode this using m bits. The best error that we can achieve is $\frac{B}{2^{m-1}}$, by performing a binary search on the range $[-B, B]$ for $g_i^{(j)}$, sending one bit per level of the binary search. However, this scheme is not collaborative. To obtain a collaborative scheme, for some permutation ρ on the set of clients $[m]$, the ith client can perform binary search until level $\rho^{(i)}$ and sends its decision at level $\rho^{(i)}$. In this case, each client sends only 1 bit per coordinate. To decode $\tilde{g}^{(j)}$, we take a weighted sum of the signs obtained from different clients weighed by their coefficients $\frac{B}{2^{\rho^{(i)}-1}}$. This is the core subroutine (Algorithm 2). The full compression scheme for d coordinates applies this coordinate-wise in Algorithm 3. Note that the clients and the server should share the permutation ρ before encoding and decoding, which need not change over different instantiations of the mean estimation problem. To understand the core idea of the scheme, consider the case when all vectors $g_i = g$. Then, sending a different level from a different client is equivalent to doing a full binary search to quantize g. As long as g_is are close to g, we hope that this scheme should give us a good estimate of g. Let $\tilde{b}_{i,k}^{(j)}$ denotes the encoding of $g_i^{(j)}$ at level k $\forall i, k \in [m], j \in [d]$. Therefore, if we were to use binary search to encode $g_i^{(j)}$, then $\sum_{k=1}^{m} \frac{B\tilde{b}_{i,k}^{(j)}}{2^{k-1}}$ would be it's decoded value.

Theorem 2 (HadamardMultiDim Error)**.** *Under Assumptions 1, the estimation error for Algorithm 3 is*

$$\mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g} - g||_\infty] \leq \frac{B}{2^{m-1}} + \min\{\Delta_{\text{Hadamard}}, \Delta_{\infty,\max}\}, \tag{5}$$

where $\Delta_{\text{Hadamard}} \equiv \max_{r\in[d]} \sqrt{\frac{1}{m^2} \sum\sum_{1\leq i\neq j\leq m} \sum_{k=1}^{m} \left(\frac{B(\tilde{b}_{i,k}^{(r)} - \tilde{b}_{j,k}^{(r)})}{2^{k-1}}\right)^2}$ *, and* $\Delta_{\infty,\max} \equiv \max_{r\in[d],i\in[m]} |g_i^{(r)} - g^{(r)}|$.

As we allow full collaboration between clients, in the worst case, we might have to incur a cost $\Delta_{\infty,\max}$ which is the worst case dissimilarity among clients. However, if client vectors are close, we might end up paying a much lower cost of Δ_{Hadamard}.

3.2 Sparse regression coding

In this part, we extend the coordinate-wise guarantee of the HadamardMultiDim to ℓ_2 error between d-dimensional vectors of bounded ℓ_2-norm.

Assumption 2 (Norm Ball)**.** $||g_i||_2 \leq B, \forall i \in [m]$.

To extend the idea of binary search and full collaboration from HadmardMultiDim, we first need a compression scheme which performs binary search on d dimensional vectors with ℓ_2 error guarantees. Sparse Regression codes (Venkataramanan et al., 2014b;a), which are known to achieve the optimal rate-distortion tradeoff

for a Gaussian source, fit our requirements. Let $A \in \mathbb{R}^{mL \times d}$ be a matrix for some parameter $L > 0$, where A_k denotes the kth row of A. To compress a single vector g by Sparse Regression codes with the matrix A, we use $m \log L$ bits to find the closest vector to g in the first L rows of A; say the index of this vector is $\tilde{b}_1$. Similar to binary search, we subtract $c_1 A_{\tilde{b}_1}$ from g, where c_1 is given in (9) to obtain an updated g. We repeat the process using the next set of L rows. Here, each set of L rows corresponds to a single level of binary search, with the coefficients c_i obtained from Eq (9) having a decaying exponent. Although (Venkataramanan et al., 2014b) requires A to be a random matrix with iid Gaussian entries, if A satisfies the following definition, then we can apply the theoretical guarantees of (Venkataramanan et al., 2014b).

Definition 1 ($(\delta_1, \delta_2, \Gamma)$-Cover)**.** *A matrix $A \in \mathbb{R}^{mL \times d}$ is a $(\delta_1, \delta_2, \Gamma)$-Cover for $\delta_1, \delta_2 \in (0, 1)$ and $\Gamma > 0$, if for each $k \in [m]$, $\exists \epsilon_k, \gamma_k, \Gamma_k \in \mathbb{R}$, such that*

$$\max_{j \in [(k-1)L+1, kL]} ||A_j||_2 \leq \sqrt{d}(1 + \gamma_k); \quad \max_{j \neq j' \in [(k-1)L+1, kL]} ||A_j - A_{j'}||_2 \leq \Gamma_k$$
$$\forall v \in \mathbb{S}^{d-1}, \max_{j \in [(k-1)L+1, kL]} \langle v, A_j \rangle \geq \sqrt{2 \log L}(1 + \epsilon_k),$$
$$\sum_{k=1}^{m} \frac{|\gamma_k|}{m} \leq \delta_1, \quad \sum_{k=1}^{m} \frac{|\epsilon_k|}{m} \leq \delta_2, \text{ and } \max_{k \in [m]} \Gamma_k \leq \Gamma.$$

By the above definition, for each $k \in [m]$, the L consecutive rows of A, which correspond to $\{A_j, j \in [(k-1)L + 1, kL]\}$, form a ε-net (Vershynin, 2018, Chapter 4.2) of $\mathbb{S}^{d-1}$ with $\varepsilon = \sqrt{2(1 - \sqrt{\frac{2 \log L}{d}})}$. By carefully selecting the parameters in the proof of (Venkataramanan et al., 2014b, Theorem 1), we can show that this scheme obtains ℓ_2 error $B \exp(-m)$ as long as A is a $(\delta_1, \delta_2, \Gamma)$-cover. Note that the matrix A is deterministically constructed with m sets of ε-nets. The choice of Gaussian A in (Venkataramanan et al., 2014b) can now be understood from the fact that such random Gaussian matrices are in fact ε-nets. The following remark summarizes this.

Remark 1 (Gaussian A)**.** *If each entry of the matrix A is sampled iid from $\mathcal{N}(0, 1)$ then, for any $\delta_1, \delta_2, \delta_3 \in [0, 1)$ and $\Gamma^2 = 2d + 4\sqrt{d \log(\frac{mL^2}{\delta_3})} + 4 \log(\frac{mL^2}{\delta_3})$, the matrix A is a $(\delta_1, \delta_2, \Gamma)$-Cover with probability $1 - \delta_3 - 2m^2 L e^{-\frac{d\delta_1^2}{8}} - m \left(\frac{L^{2\delta_2}}{\log L}\right)^{-m}$.*

Note that the expressions for δ_1, δ_2 can be obtained from (Venkataramanan et al., 2014b) and the proof for Γ is provided in Appendix B.3. For $d = \Omega(\log m)$, the probability of error in the above Remark can be made arbitrarily close to 1 for large m.

We now extend the scheme based on sparse regression codes in (Venkataramanan et al., 2014b) to our setting of distributed mean estimation, similar to HadamardMultiDim. Each client $i \in [m]$ encodes at level $\rho^{(i)}$ where ρ is a permutation on $[m]$ and the server computes the weighted sum of the encodings from each client with corresponding coefficients $c_{\rho^{(i)}}$. Here, $\tilde{b}_{i,k}$ denotes the encoding of g_i at level k $\forall i, k \in [m]$. If we were to use Sparse Regression codes to encode only g_i then $\sum_{k=1}^{m} c_k A_{(k-1)L + \tilde{b}_{i,k}}$ would be it's decoded value. The full scheme in described in Algorithm 4 and we compute it's ℓ_2 estimation error in the next theorem. In this scheme, each client only needs to send a single index $\tilde{b}_{i,\rho^{(i)}}$. As ρ is known to the server, this index is in the range $[L]$ and each client uses only $\log L$ bits for communication.

Theorem 3 (SparseReg Error)**.** *Under Assumption 2, the following statements hold for $\tilde{g}$, the output of Algorithm 4 for any matrix $A \in \mathbb{R}^{mL \times d}$.*

$$\mathbb{E}_{\rho \sim \Pi_m}[||g - \tilde{g}||_2^2] = ||\bar{g} - g||_2^2 + \Delta_{\text{reg}}, \tag{6}$$
$$\text{where, } \Delta_{\text{reg}} \equiv \frac{1}{m^2} \sum_{i,j \in [m], i \neq j} \sum \sum_{k=1}^{m} c_k^2 ||A_{(k-1)L + \tilde{b}_{i,k}} - A_{(k-1)L + \tilde{b}_{j,k}}||_2^2 \text{ and } \bar{g} \equiv \frac{1}{m} \sum_{i=1}^{m} \sum_{k=1}^{m} c_k A_{(k-1)L + \tilde{b}_{i,k}}.$$

In addition, if the matrix A is a $(\delta_1, \delta_2, \Gamma)$-Cover, then,

$$\mathbb{E}_{\rho\sim\Pi_m}[||g-\tilde{g}||_2^2] \leq \Lambda + \Delta_{2,\max}, \quad \textit{and,} \quad ||\bar{g}-g||_2^2 \leq \Lambda, \tag{7}$$

$$\textit{where, } \Lambda \equiv B^2\left(1+\frac{10\log L}{d}e^{\frac{m\log L}{d}}(\delta_1+\delta_2)\right)^2\left(1-\frac{2\log L}{d}\right)^m.$$

$$\Delta_{\text{reg}} \leq \frac{2B^2\Gamma^2\log L}{d^2m^2}\sum\sum_{i,j\in[m],i\neq j}\sum_{k=1}^{m}\left(1-\frac{2\log L}{d}\right)^{k-1}\mathbf{1}(\tilde{b}_{i,k}\neq\tilde{b}_{j,k}). \tag{8}$$

Remark 2 (Growth of Λ with m)**.** *By expanding the squares, we can see that Λ can be bounded by 3 terms.*

$$\Lambda \leq B^2e^{-\frac{2m\log L}{d}} + \frac{20B^2\log L}{d}(\delta_1+\delta_2)e^{-\frac{m\log L}{d}} + \frac{100B^2\log^2 L}{d^2}(\delta_1+\delta_2)^2.$$

The first and the second terms decrease exponentially with m. The third term is a small constant independent of m, depending on $\delta_1,\delta_2 \ll 1$ that can be made arbitrarily small.

The proof is provided in Appendix B.2. Similar to HadmardMultiDim, the first term in the estimation error, Λ, has an exponential dependence on m and is obtained from the existing results of Sparse Regression Codes from (Venkataramanan et al., 2014b). In terms of ℓ_2 error this dependence on m is better than all the prior methods. Note that if our matrix A is a $(0,0,\Gamma)$-Cover, i.e., $\delta_1=\delta_2=0$, then Λ decreases exponentially with m. As mentioned in Remark 1, we can can achieve very small (but nonzero) values for δ_1,δ_2 with Gaussian matrices.

By combining Eqs (6) and (7), we can bound the squared ℓ_2 estimation error by $\Lambda+\min\{\Delta_{\text{reg}},\Delta_{2,\max}\}$. Note that this error resembles Theorem 2, as it contains an exponentially decreasing term of m along with a minimum of two terms, one of which is the maximum deviation between any client and the other($\Delta_{2,\max}$) and the other is the average difference of encoded values across all pairs of clients ((Δ_{reg})). Further, even the dissimilarity term Δ_{reg} has a similar structure to Δ_{Hadamard} as it is the pairwise difference between encodings of two different vectors at all levels, where the difference at level k is proportional to e^{-k}. If the client vectors are not close to each other, we might incur the worst possible error $\Delta_{2,\max}$, but if they are close, we will pay an average price in terms of Δ_{reg}.

The upper bound on Δ_{reg} in Eq (8) depends on A only via the encoded values $\{\tilde{b}_{i,k}\}_{i,k\in[m]}$. For this upper bound, the corresponding terms for each level, here denoted by k, decrease geometrically with k. The term $\frac{2B^2\Gamma^2\log L}{d^2m^2}$, ensures that Δ_{reg} is of the correct scale as in the worst case when all encodings are different for all pairs of clients, the summation gives us $\Theta(\frac{d}{2\log L})$ which results in an overall Δ_{reg} of the scale of B^2. From (Venkataramanan et al., 2014a), for each $k\in[m]$, all encodings upto and including the kth encoding, i.e., $\{\tilde{b}_{i,k'}\}_{k'\leq k}$, form an estimate of g_i that has error scaling as $Be^{-\frac{\log L(k-1)}{d}}$, which is similar to the error scaling of HadamardMultiDim for a single dimension.

While both the HadmardMultiDim and SparseReg schemes achieve very low communication rate, that comes at the price of $O(m)$ computing in the `Encode` step. This higher cost in computing is to be expected when one wants to exploit the full potential of collaborative compression (e.g., (Jiang et al., 2023), where the `Decode` step takes $O(m^2)$ time).

3.3 Benefits of HadmardMultiDim and SparseReg

We now provide a example to show that for practical scenarios, the error terms Δ_{reg} and $\Delta_{\text{H}adamard}$ are much smaller than their worst case values $\Delta_{\infty,\max}$ and $\Delta_{2,\max}$ and have similar behavior as their average values, $\Delta_2 \equiv \frac{1}{m}\sum_{i=1}^m||g_i-g||_2^2$ and $\Delta_\infty\Delta_\infty \equiv \max_{r\in[d]}\frac{1}{m}\sum_{i=1}^m|g_i^{(r)}-g^{(r)}|$ respectively. Consider the scenario of Theorem 2 (ℓ_∞ error) and set $d=1$. Assume that the first c vectors are g_1' and the remaining $m-c$ vectors are g_2', for some constant $c\ll m$. In this case, $\Delta_{\infty,\max}=(1-\frac{c}{m})|g_1'-g_2'|\approx|g_1'-g_2'|$, while $\Delta_\infty\approx\frac{2c}{m}|g_1'-g_2'|$. In this scenario, if the compressed values $\tilde{b}$ for g_1' and g_2' according to the HadamardMultiDim differ at $k\in\mathcal{K}\subseteq[m]$ levels , then, $\Delta_{\text{Hadamard}}\approx\sqrt{\frac{c}{m}\sum_{k\in\mathcal{K}}(B/2^{k-1})^2}\leq\sqrt{\frac{cB^2}{m}\sum_{k=k^\star}^{\infty}2^{-2(k-1)}}\leq\sqrt{\frac{4c}{3m}}\frac{B}{2^{k^\star-1}}$, where

$k^\star = \min_{k\in\mathcal{K}} k$ and we upper bound the sum over the set $\mathcal{K}$ by the sum of a geometric series. **As Δ_{Hadamard} averages over all machines, it decreases with m similar to Δ_2 and should be much smaller than $\Delta_{\infty,\max}$.** The only case when it is not smaller than $\Delta_{\infty,\max}$ is when g_1' and g_2' are very close, so that $\Delta_{\infty,\max} = \mathcal{O}(\sqrt{m^{-1}})$, but the first level where they differ ($\min_{k\in\mathcal{K}} k$) is very small. One such example is when the quantized values of g_1' in the set $\mathcal{K}$ sorted by the levels in increasing order are $(+1,-1,-1,-1)$ and that of g_2' are $(-1,+1,+1,+1)$. As the vectors are extremely close in this case, the estimation error with $\Delta_{\infty,\max}$ is not very large. Further, if we assume a distributional assumption on the vectors g_i, similar to how we generate Figure 2b, obtaining vectors where $\Delta_{\text{Hadamard}} > \Delta_{\infty,\max}$, happens with low probability.

Algorithm 4 SparseReg

`Init()`
Clients and server share $A \in \mathbb{R}^{mL\times d}$, and a permutation $\rho \sim \Pi_m$.
`Encode(`g_i`)`
$g_i' \leftarrow g_i$
for $j \in [\rho^{(i)}]$ **do**
  $\tilde{b}_{i,j} \leftarrow \arg\max_{r\in[L]} \langle A_{(j-1)L+r}, g_i'\rangle$
  $g_i' \leftarrow g_i' - c_j A_{(j-1)L+\tilde{b}_{i,j}}$
end for
$\tilde{b}_i \leftarrow \tilde{b}_{i,\rho^{(i)}}$
return $\tilde{b}_i$
`Decode(`$\{\tilde{b}_i\}_{i\in[m]}$`)`
$\tilde{g} \leftarrow \sum_{i\in[m]} c_{\rho^{(i)}} A_{(\rho^{(i)}-1)L+\tilde{b}_i}$

$$c_i = B\sqrt{\frac{2\log L}{d^2}\left(1-\frac{2\log L}{d}\right)^{i-1}} \tag{9}$$

Algorithm 5 OneBit

`Init()`
Clients and server share unit vectors $\{z_i\}_{i\in[m]}$.
`Encode(`g_i`)`
$\tilde{b}_i \leftarrow \text{sign}(\langle g_i, z_i\rangle)$
return $\tilde{b}_i$
`Decode(`$\{\tilde{b}_i\}_{i\in[m]}$`)`
$g' \leftarrow \begin{cases} \text{(Shen, 2023, Algorithm 1) (Tech. I)} \\ \frac{1}{m}\sum_{i=1}^m z_i\tilde{b}_i \text{ (Tech. II)} \end{cases}$
$\tilde{g} \leftarrow g'/||g'||_2$

Using this example, we can compare the error of our proposed schemes to baselines mentioned in Table 2. Assume a modification of the above example where each vector is d-dimensional, and all of it's coordinates are either $g_1' \nparallel_d$ or g_2'. Consider any ℓ_2 compressor whose error is either proportional to $\Psi\tilde{B}^2$ or $\Psi\Delta_2$ and it sends ψ bits/client for some constants $\psi, \Psi > 0$. The ℓ_2 error is defined as $\mathbb{E}[||\tilde{g}-g||_2^2]$ and the ℓ_∞ error is defined as $\mathbb{E}[||\tilde{g}-g||_\infty]$, therefore the corresponding ℓ_∞ error of these compressors is $\sqrt{\Psi}\tilde{B}$ or $\sqrt{\Psi}\Delta_2$. Now, consider the example that we just presented with $d > 1$ and all coordinates being equal for each vector. Therefore, $\Delta_2 \approx \frac{cd}{m}|g_2' - g_1'|^2$, and plugging this in, the ℓ_2 error of the schemes is $\sqrt{\Psi}\tilde{B}$ or $\sqrt{\Psi\frac{cd}{m}}|g_2'-g_1'|$. HadamardMultiDim sends d bits/client; therefore, to compare with any of these schemes, we set $\psi = d$. For RandK, this would mean setting $K = \frac{d}{32+\log d}$. Now, if $|g_1'|, |g_2'| \approx B$ but $|g_2'-g_1'| \ll B$, then $\tilde{B} \approx \sqrt{d}B$. Using these approximations, the error of RandK is $\sqrt{(32+\log d)d}B$ as $\Psi = \frac{d}{K} = 32+\log d$. This is much larger than the ℓ_∞ error of HadamardMultiDim, as the first term is $B\cdot 2^{m-1}$ and the second term $\Delta_{\text{Hadamard}} \approx \sqrt{\frac{c}{m}}|g_2'-g_1'|$. A similar argument holds for all independent compression schemes, as their ℓ_∞ error scales as $\tilde{B}$, which in the worst case is $\sqrt{d}B$.

For compressors whose error scales as $\Psi\Delta_2$ (PermK, RandKSpatial, RandKSpatialProj), we obtain the same number of bits/client as HadamardMultiDim scheme by setting $K = \frac{d}{32+\log d}$. Consider RandKSpatialProj, where $\Psi = \frac{32+\log d}{m}$, and the error for our example is $\sqrt{c\frac{(32+\log d)d}{m^2}}|g_2'-g_1'|$. As long as $d > m$, this error is larger than Δ_{Hadamard} by constant terms. A similar argument holds for RandKSpatial and PermK. Additionally, note that the theoretical guarantees for RandKSpatial and RandKSpatialProj do not hold if the correlation is unknown. Without this information, the heuristics they use do not result in theoretical guarantees, and their error might be as bad as RandK. CorrelatedSRQ achieves the lower bound for collaborative compressors for $d = 1$, and is based on a coordinate-wise scheme, hence the Δ_∞ in its error guarantees. However, for $d \gg 1$, its error scales poorly. For the example described above, $||g_i||_2 \le \sqrt{d}B$, therefore, the ℓ_∞ error for CorrelatedSRQ is $\sqrt{\frac{1}{m}\min\{\frac{d\Delta_\infty^d B}{K}, \frac{d^3B^2}{K^2}\}}$. Even for $K = 2$, correlated SRQ requires double the number of bits/client as HadamardMultiDim. The first term of HadamardMultiDim is $B\cdot 2^{m-1}$, which is much smaller than any of these terms, while $\Delta_{\text{Hadamard}} \approx \sqrt{\frac{m}{c}}\Delta_\infty$ for our example. Therefore, as long

as $\left(\frac{m^2K}{cdB}\right)^{1/(2d-1)} < \Delta_\infty < \frac{\sqrt{c}dB}{mK}$, Δ_{Hadamard} is smaller than ℓ_∞ error of CorrelatedSRQ. The size of this interval for Δ_∞ increases as d increases.

With the above example and analysis, we have specified the exact scenarios when HadamardMultiDim outperforms baselines, and this can be easily extended to SparseReg.

4 One-bit Schemes

In this section, our vectors are assumed to belong on the unit sphere $\mathbb{S}^{d-1}$. Further, our goal is to recover the unit vector in the direction of the average vector $g = (\frac{1}{m}\sum_{i\in[m]} g_i)/||\frac{1}{m}\sum_{i\in[m]} g_i||_2$.

Assumption 3 (Unit vectors)**.** $g_i \in \mathbb{S}^{d-1}, \forall i \in [m]$.

Consider the collaborative compressor where each client has sample $z_i \sim \text{Unif}(\mathbb{S}^{d-1})$ (which are also available to the server apriori). Client i sends the single bit $\tilde{b}_i = \text{sign}(\langle g_i, z_i\rangle)$ to the server. To recover g, consider the trivial case when all vectors g_is were equal. Then, each $\tilde{b}_i = \text{sign}(\langle g, z_i\rangle)$, and to recover g, the server needs to learn the halfspace corresponding to g from a set of m labeled datapoints. Applying the same method to when g_is are not all the same, we can estimate g by solving the following optimization problem.

$$\min_{\tilde{g}\in\mathbb{S}^{d-1}} \frac{1}{m}\mathbf{1}(\tilde{b_i} \neq \text{sign}(\langle z_i, \tilde{g}\rangle)). \tag{10}$$

Here, $\mathbf{1}(\cdot)$ denotes the indicator function. We can intuitively view (10) as a halfspace learning problem with a groundtruth g, but in the presence of noise, as $g_i \neq g$. Learning halfspaces in the presence of noise is hard in general (Guruswami & Raghavendra, 2006). In our setting, if we sample z_i from the intersection of the halfspaces with normal vectors g and g_i, then the label is $\text{sign}(\langle g, z_i\rangle)$, otherwise, it is $-\text{sign}(\langle g, z_i\rangle)$. We can consider this to be under the malicious noise model, wherein a fraction of datapoints are corrupted.

Lemma 1 (Malicious Noise)**.** *If $z_i \sim \text{Unif}(\mathbb{S}^{d-1})$ and $\tilde{b_i} = \text{sign}(\langle z_i, g_i\rangle)$, $\forall i \in [m]$, then, with probability $1 - \mathcal{O}(\exp(-m\Delta_{\text{corr}}))$, ζ, the fraction of the set of datapoints $\{(z_i, \tilde{b_i})\}_{i\in[m]}$ satisfying $\text{sign}(\langle z_i, g_i\rangle) \neq \text{sign}(\langle g, z_i\rangle)$ is equal to $\Theta(\Delta_{\text{corr}})$, where $\Delta_{\text{corr}} \triangleq \frac{1}{m\pi}\sum_{i=1}^m \arccos(\langle g_i, g\rangle)$.*

The proof of the lemma is provided in Appendix C.1. Our methods will use Δ_{corr} to measure the deviation between clients. For small Δ_{corr}, we obtain better performance. If $\langle g, g_i\rangle \geq 0, \forall i \in [m]$, then

$$\cos(\pi\Delta_{\text{corr}}) \geq \sqrt{\frac{1}{m} + \frac{2}{m^2}\sum\sum_{1\leq i<j\leq m}\langle g_i, g_j\rangle}. \tag{11}$$

As long as the corruption level, $\zeta < \frac{1}{2}$, we can hope to recover the halfspace g. We provide two techniques – Techniques I and II, to recover g, thus yielding two corresponding `Decode` procedures.

The first decoding procedure (Technique I) is a linear time algorithm for halfspace learning in the presence of malicious noise (Shen, 2023, Theorem 3) that provides optimal sample complexity and noise tolerance.

Theorem 4 (Error of Technique I)**.** *If ζ defined in Lemma 1 is less than $\frac{1}{2}$, after running Algorithm 5 with Technique I, with probability $1 - \delta - \mathcal{O}(\exp(-m\Delta_{\text{corr}}))$, we obtain a hyperplane $\tilde{g}$ such that, $\langle \tilde{g}, g\rangle \geq \cos(\pi(\Delta_{\text{corr}} + \frac{d}{m}\text{polylog}(d, \frac{1}{m}, \frac{1}{\delta}))$.*

The algorithm itself is fairly complicated. It assigns weights to different points based on how likely they are to be corrupted. The algorithm proceeds in stages, wherein each stage decreases the weights of the corrupted points and solves the weighted version of (10). The key technique is to use matrix multiplicative weights update (MMWU) (Arora et al., 2012) to yield linear time implementation of both these steps, instead of (Awasthi et al., 2017) which used polynomial time linear programs for this purpose.

Technique II is the simple average algorithm of (Servedio, 2002), which obtains suboptimal error guarantees for halfspace learning.

Theorem 5 (Error of Technique II)**.** *If ζ defined in in Lemma 1 is less than $\frac{1}{2}$, after running Algorithm 5 with Technique II, with probability $1 - \delta - \mathcal{O}(\exp(-m\Delta_{\text{corr}}))$, we obtain a hyperplane $\tilde{g}$ such that, $\langle \tilde{g}, g \rangle \geq \cos(\pi(\sqrt{d}\Delta_{\text{corr}} + \frac{d}{\sqrt{m}}\sqrt{\log(\frac{d}{\delta})}))$.*

The performance of both techniques improves with decrease in Δ_{corr}. Since we have only m bits to infer a d-dimensional vector, we require $m > d$, with Technique II requiring $m > d^2$. If we send t bits per client in OneBit, then the number of samples for the halfspace learning is mt, thus obtaining the guarantee in Table 1. The main benefit of OneBit schemes is their extreme communication efficiency. Existing quantization and sparsification schemes require sending at least $\log K$ or $\log d$, where K is the number of quantization levels. The proofs for this section are provided in Appendix C.

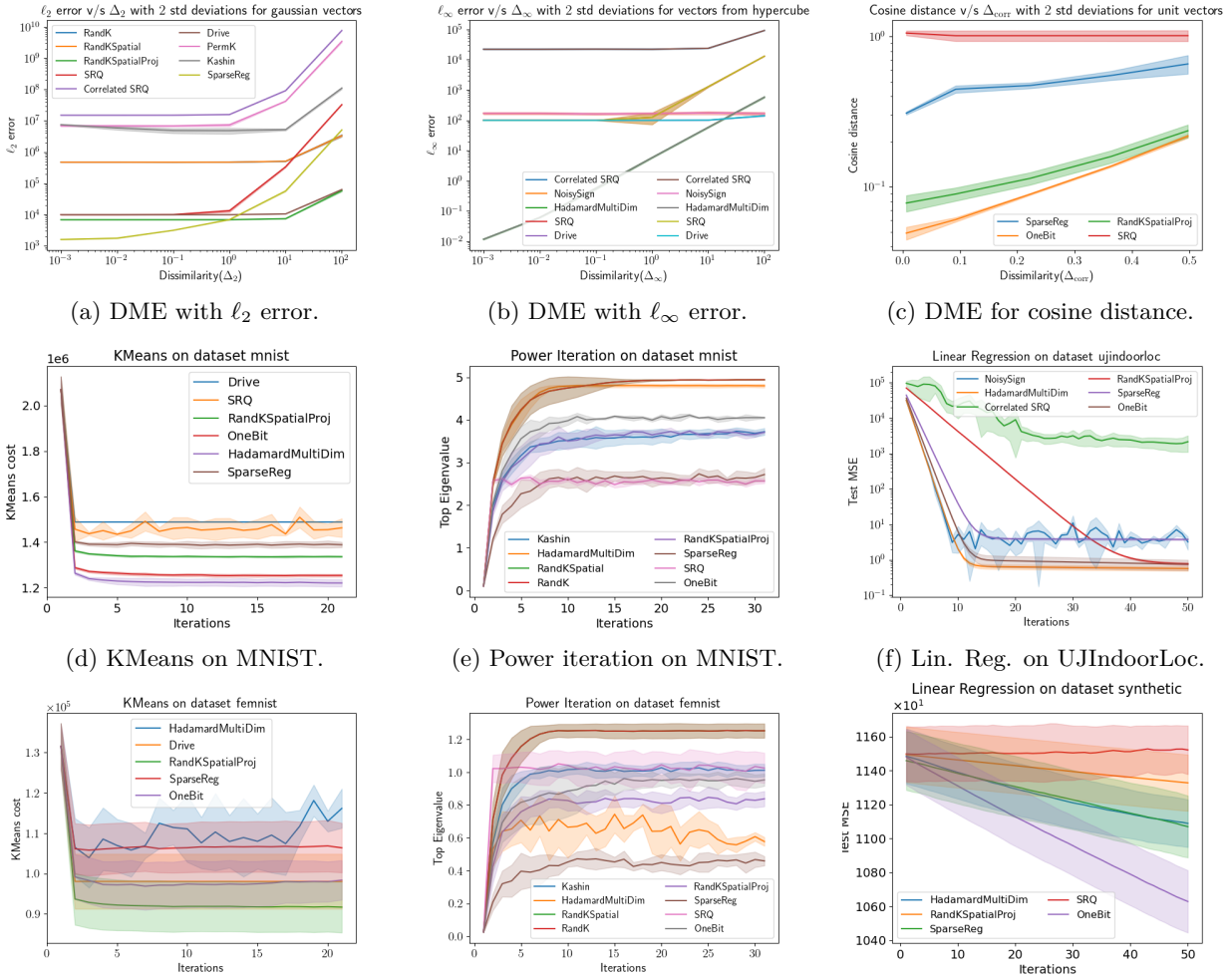
Comparison to ℓ_2 compressor. Note that, we can use compressor for ℓ_2 error to first decode the mean and then normalize it to obtain its unit vector. If such a scheme uses t bits and has ℓ_2 error either $\Lambda\Delta_2$ or $\Lambda\tilde{B}^2$ then its cosine similarity $\frac{\langle g, \tilde{g} \rangle}{||g'||_2||\tilde{g}||_2} \geq 1 - \frac{\Lambda}{2||g'||_2^2}$ for $||g'||_2 \approx ||\tilde{g}||_2$, where $g' = \frac{1}{m}\sum_{i=1}^{m} g_i$ and $\tilde{g}$ is the estimate of g'. To compare this with OneBit Technique I, we send λ bits per client to obtain the same communication budget. The cosine similarity of this scheme is $\cos(\pi(\Delta_{\text{corr}} + \frac{d}{tm}))$. We can lower bound this similarity by $1 - 2\pi^2\Delta_{\text{corr}}^2 + 2\pi^2\frac{d^2}{m^2t^2}$ as $\cos(x) \geq 1 - \frac{x^2}{2}$. Comparing this cosine similarity with that obtained for ℓ_2-compressor, as long as $2\pi^2\Delta_{\text{corr}}^2 + 2\pi^2\frac{d^2}{m^2\beta^2} < \Lambda$, OneBit Technique I performs better. For any sparsification scheme sending K coordinates, Λ is at least $\frac{d}{mK}$. If we set $t = 32K + K\log d$, OneBit Technique I outerperforms the sparsification scheme as long as Δ_{corr} is small.

5 Experiments

Setup. To compare the performance of our proposed algorithms, we perform DME for three different distributions which correspond to the three error metrics covered by our schemes – ℓ_2, ℓ_∞ and cosine distance. Then, we run our algorithms as the DME subroutine for four different downstream distributed learning tasks – KMeans, power iteration, linear regression and logistic regression. KMeans and power iteration are run on MNIST (LeCun & Cortes, 2010) and FEMNIST (Caldas et al., 2018) datasets, and we report the KMeans cost and top eigenvalue as the metrics. For linear regression, we run gradient descent on UJIndoorLoc (Torres-Sospedra et al., 2014) and a Synthetic mixture of regressions dataset, with low dissimilarity between the mixture components, and report the test MSE. For logistic regression, we run gradient descent on the HAR (Reyes-Ortiz et al., 2012) dataset and report the training loss and test accuracy for binary classification. We compare against all baselines in Table 2 for 5 random seeds and report the methods which perform the best. For DME experiemnts, the shaded corresponds to 2 standard deviations around the mean, while for all downstream tasks, it corresponds to 1 standard deviation. The results for KMeans, power iteration and linear regression are reported in Fig 2, while the results for logistic regression are reported in Fig 3. Additional details for our experimental setup are deferred to Appendix D.

Results. *Distributed Mean Estimation.* From Fig 2a and 2b, HadamardMultiDim and SparseReg, whose error is optimal in m, obtain the best performance in terms of ℓ_∞ and ℓ_2 error for low dissimilarity. Especially, for HadamardMultiDim in Fig 2b, the gap in ℓ_∞ error to next best scheme is very large. NoisySign obtains competitive performance to other baselines as we use a large σ. The performance of OneBit for cosine distance metric (Fig 2c) shows that compressors with ℓ_2 error guarantees perform poorly in terms of cosine distance. For all collaborative compression schemes, including our proposed schemes, performance degrades as dissmilarity increases. From Fig 2a and 2b, the rate of this decrease is more severe for SparseReg than HadamardMultiDim. For large dissimilarity, HadamardMultiDim and SparseReg can perform worse than certain baselines.

KMeans and Power iteration. For MNIST dataset, where dissimilarity is low, HadamardMultiDim performs best for KMeans and close to the best baseline for power iteration (Fig 2d and 2e). Most of our collaborative compression schemes do not perform as well as RandK on FEMNIST, due to higher client dissimilarity. OneBit is very communication-efficient, so running it for the same communication budget as our baselines ensures that it still remains competitive for KMeans(Fig 2g). By increase in confidence intervals for the same baselines from MNIST to FEMNIST, we can see that FEMNIST dataset is much more heterogeneous.



(a) DME with ℓ_2 error. (b) DME with ℓ_∞ error. (c) DME for cosine distance.

(d) KMeans on MNIST. (e) Power iteration on MNIST. (f) Lin. Reg. on UJIndoorLoc.

(g) KMeans on FEMNIST. (h) Power iteration on FEMNIST. (i) Lin. Reg. on Synthetic.

Figure 2: Performance of DME(Distributed Mean Estimation), KMeans, Power iteration and linear regression for the same communication budget. For each experiment, we report the best compressors. Lin. Reg. refer to Linear Regression. For power iteration, higher top eigenvalue is better. For all other experiments, we report the error, so lower is better.

Linear Regression. From Fig 2f and2i, all collaborative compressors perform better than independent compressors as UJIndoorLoc and synthetic datasets have low dissimilarity among clients as compared to FEMNIST. Our schemes can take full advantage of this low dissimilarity, so HadamardMultiDim and SparseReg outperform baselines on both datasets. As the Synthetic dataset has lower dissimilarity than UJIndoorLoc, even the NoisySign performs better than other baselines. Further, OneBit obtains best performance.

Logistic Regression. From Fig 3, our compressors, Onebit, Sparsereg, and HadamardMultDim, are the best, second, and fourth best compressors, respectively, in terms of both training loss and test accuracy. Further, among the top 4 best-performing schemes, only one baseline, RandKSpatialProj, comes in third. This shows the benefit of using collaborative compressors.

Stability of Our Algorithms. Among all our proposed algorithms, OneBit and SparseReg are the most stable with respect to different random seeds followed by HadamardMultiDim and NoisySign. The most stable algorithms among all baselines seems to be RandKSpatialProj. OneBit and Sparsereg are slightly less stable

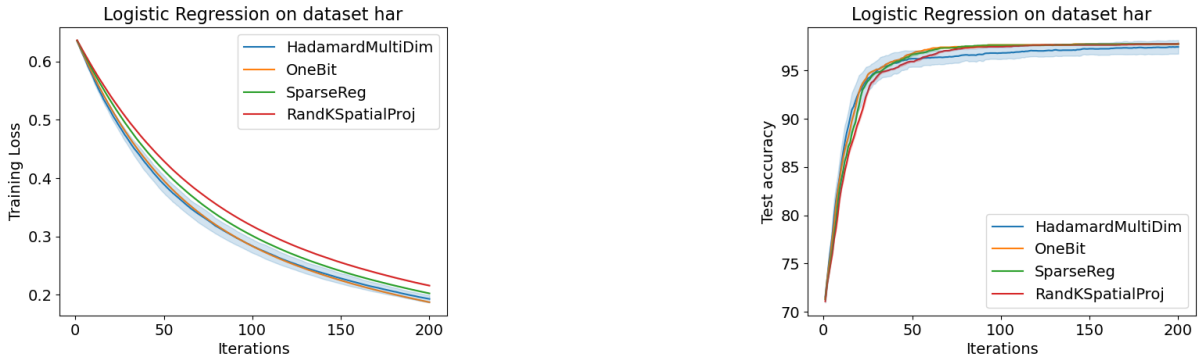


Figure 3: Performance of compressors for Logistic regression on HAR (Reyes-Ortiz et al., 2012) dataset. Left: Training Logistic loss, Right : Test Accuracy.

than it. However, all our algorithms except NoisySign are much more stable than most other baselines like SRQ, Correlated SRQ and RandK.

6 Conclusion

We proposed four communication-efficient collaborative compression schemes to obtain error guarantees in ℓ_2-error (SparseReg), ℓ_∞-error (NoisySign, HadamardMultiDim) and cosine distance (OneBitAvg). The estimation error of our schemes improves with number of clients, and degrades with increasing dissimilarity between clients. Our schemes are biased and our dissimilarity metrics (Δ_{reg}, Δ_{Hadamard}) depend on the quantization levels. However, these drawbacks can be removed by using existing techniques for converting biased compressors to unbiased ones (Beznosikov et al., 2022), and adding noise before quantization (Tang et al., 2023; Chzhen & Schechtman, 2023). Schemes such as *error feedback* (Karimireddy et al., 2019) reduces the error of independent compressors, and it will be interesting to check if it works for our collaborative compressors.

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A Proofs for Section 2

A.1 Proof of Theorem 1

As all operations are coordinate-wise, we restrict our focus to only a single dimension $j \in [d]$.

$$\mathbb{E}_{\xi_i}[\tilde{b}_i^{(j)}] = \Phi_\sigma(g_i^{(j)}),\ \forall i \in [m]$$

Note that $\Phi_\sigma(t) = \mathrm{erf}(\frac{t}{\sqrt{2}\sigma})$ and $\Phi_\sigma^{-1}(t) = \sqrt{2}\sigma\,\mathrm{erf}^{-1}(t)$. By Hoeffding's inequality for bounded random variables we have,

$$\Pr[|\frac{1}{m}\sum_{i=1}^{m}(\tilde{b_i}^{(j)} - \Phi_\sigma(g_i^{(j)}))| \geq t] \leq 2e^{-\frac{mt^2}{2}}$$

If we set $t = \sqrt{\frac{2c\log(m)}{m}}$, for some $c > 0$ in the above inequality, then with probability $1 - 2m^{-c}$, we have,

$$|\frac{1}{m}\sum_{i=1}^{m}(\tilde{b_i}^{(j)} - \Phi_\sigma(g_i^{(j)}))| \leq t$$

We can represent $\frac{1}{m}\sum_{i=1}^{m}\tilde{b}_i = \Phi_\sigma(\tilde{g})$, as Φ_σ is an invertible function. To find the difference between $\tilde{g}$ and g, we find the difference $\Phi_\sigma(\tilde{g}) - \Phi_\sigma(g)$. With probability $1 - 2m^{-c}$, we have,

$$|\Phi_\sigma(\tilde{g}^{(j)}) - \Phi_\sigma(g^{(j)})| \leq \frac{1}{m}\sum_{i=1}^{m}|\Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})| + t$$

To remove the terms of Φ_σ, we can apply the function Φ_σ^{-1} on $\tilde{g}^{(j)}$. As Φ_σ^{-1} is not Lipschitz, we need to perform its Taylor's expansion around $\Phi_\sigma(g^{(j)})$ to account for the linear terms in the error. If $\Delta_\Phi = \frac{1}{m}\sum_{i=1}^{m}|\Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})|$, then we obtain,

$$|\tilde{g}^{(j)} - g^{(j)}| \leq \max_{u \in [\Phi_\sigma(g^{(j)}) - \Delta_\Phi - t, \Phi_\sigma(g^{(j)}) + \Delta_\Phi + t]} |(\Phi_\sigma^{-1})'(u)|(\Delta_\Phi + t) \tag{12}$$

We now obtain an appropriate upper bound on $(\Phi_\sigma^{-1})'(u)$ as we do not have a closed-form expression for it. We will use the properties of erf to obtain a suitable bound. First, note that Φ_σ and Φ_σ^{-1} are both odd functions, therefore, $|\Phi^{-1}(u)| = |\Phi^{-1}(|u|)|$, so we consider the bound for $u > 0$. Note that $(\Phi^{-1})'(u) = \frac{1}{\Phi'(\Phi^{-1}(u))}$. For $u > 0$, we have,

$$\begin{aligned}
1 - \mathrm{erf}(u) \leq & e^{-u^2} \\
\mathrm{erf}(u) \geq & 1 - e^{-u^2} \\
\mathrm{erf}^{-1}(u) \leq & \sqrt{-\log(1-u)} \\
\Phi_\sigma^{-1}(u) = \sqrt{2}\sigma\,\mathrm{erf}^{-1}(u) \leq & \sigma\sqrt{-2\log(1-u)} \\
(\Phi_\sigma^{-1})'(u) = \sqrt{\frac{\pi}{2}}e^{(\Phi_\sigma^{-1}(u))^2/(2\sigma^2)} \leq & \sqrt{\frac{\pi}{2}}e^{-2\log(1-u)/2} = \sqrt{\frac{\pi}{2}}\frac{1}{1-u}
\end{aligned}$$

For the first step, we use an upper bound on the complementary error function. For the third step, we use the fact that if $f(x) \leq g(x)$, then $f^{-1}(y) \geq g^{-1}(y)$.

Using the following upper bound in Eq (12), we obtain,

$$\begin{aligned}|\tilde{g}^{(j)} - g^{(j)}| \leq & \max_{u \in [\Phi_\sigma(g^{(j)}) - \Delta_\Phi - t, \Phi_\sigma(g^{(j)}) + \Delta_\Phi + t]} \sqrt{\frac{\pi}{2}} \frac{\Delta_\Phi + t}{1 - |u|} \\ \leq & \sqrt{\frac{\pi}{2}} \frac{\Delta_\Phi + t}{1 - \max\{|\Phi_\sigma(g^{(j)}) - \Delta_\Phi - t|, |\Phi_\sigma(g^{(j)}) + \Delta_\Phi + t|\}}\end{aligned}$$

We use $\max\{|\Phi_\sigma(g^{(j)}) - \Delta_\Phi - t|, |\Phi_\sigma(g^{(j)}) + \Delta_\Phi + t|\} \leq \Phi_\sigma(|g^{(j)}|) + \Delta_\Phi + t$, as Φ_σ is an increasing odd function.

$$|\tilde{g}^{(j)} - g^{(j)}| \leq \sqrt{\frac{\pi}{2}} \left(\left(1 - \frac{\Delta_\Phi + t}{1 - \Phi_\sigma(|g^{(j)}|)} \right)^{-1} - 1 \right)$$

We extend the bound to d dimensions by taking a union bound, yielding a probability of error $2dm^{-c}$.

A.2 Proof of Equation (3)

First, note that the term of $(1+x)^{-1} - 1 = \frac{x}{1-x} \leq \mathcal{O}(x)$ when $x \leq c$ for some constant $c \in [0, 1]$. The conditions required for this are Δ_Φ to be small and both m and $\alpha(||g||_\infty)$ to be large. Note that $\alpha(||g||_\infty) = 1 - \Phi_\sigma(||g||_\infty)$, so it is small when the ratio $\frac{||g||_\infty}{\sigma}$ is small. The bound on estimation error after this step is the following.

$$||\tilde{g} - g||_\infty = \mathcal{O}\left(\frac{\Delta_\Phi + \sqrt{\frac{\log m}{m}}}{\alpha(||g||_\infty)} \right)$$

Now, we lower bound the term $\alpha(||g||_\infty)$. Note that $\alpha(x) = 1 - \Phi_\sigma(x) = \frac{1}{2}(1 - \text{erf}(\frac{x}{\sqrt{2}\sigma}))$ for any $x \in \mathbb{R}_+$ where erf is the error function. Here, $1 - \text{erf}(x)$ is the complementary error function for any $x \geq 0$ and it is lower bounded by $\Omega(e^{-\beta' x^2})$ for $\beta' > 1$ (Chang et al., 2011). Plugging in this lower bound, we complete the proof for a $\beta = \beta'/2 > \frac{1}{2}$.

A.3 Proof of Equation (4)

The proof follows from using the triangle inequality and a Taylor's expansion for each $\Phi_\sigma(g_i^{(j)})$ around $g^{(j)}$. Note that, for some $u_i^{(j)}$ between $g^{(j)}$ and $g_i^{(j)}$, we have,

$$\begin{aligned}\Phi_\sigma(g_i^{(j)}) &= \Phi_\sigma(g^{(j)}) + \sqrt{\frac{2}{\pi}} \frac{(g^{(j)} - g_i^{(j)}) e^{-\frac{(u_i^{(j)})^2}{2\sigma^2}})}{\sigma} \\ |\Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})| &\leq \sqrt{\frac{2}{\pi}} \frac{|g^{(j)} - g_i^{(j)}|}{\sigma}\end{aligned}$$

We use the fact that $e^{-\frac{(u_i^{(j)})^2}{2\sigma^2}} \leq 1$. By using triangle inequality for any coordinate $j \in [m]$, we obtain,

$$\begin{aligned}\Delta_\Phi \leq & \max_{j \in [d]} \frac{1}{m} \sum_{i \in [m]} |\Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})| \leq \frac{1}{m} \sum_{i \in [m]} \max_{j \in [d]} |\Phi_\sigma(g_i^{(j)}) - \Phi_\sigma(g^{(j)})| \\ \leq & \sqrt{\frac{2}{\pi}} \frac{1}{m} \sum_{i \in [m]} \max_{j \in [d]} \frac{|g^{(j)} - g_i^{(j)}|}{\sigma} \leq \sqrt{\frac{2}{\pi}} \frac{1}{m} \sum_{i \in [m]} \frac{||g - g_i||_\infty}{\sigma}\end{aligned}$$

B Proofs for Section 3

B.1 Proof of Theorem 2

To prove this Lemma, we first consider a single dimension $j \in [d]$. Note that the collaborative compression scheme has precision $\frac{B}{2^{m-1}}$. Since, for each client $g_i^{(j)}$ lies in the range $[g^{(j)} - \Delta_{\max}, g^{(j)} + \Delta_{\max}]$, therefore, we must decode one of the elements which has error atmost $\Delta_{\max} + \frac{B}{2^{m-1}}$. Since this is true for each coordinate individually, therefore, it holds for the ℓ_∞ norm.

Consider a single dimension $j \in [d]$. Let $g_i^{(j)}$ be the jth coordinate of g_i and ρ_j be the permutation selected for the coordinate j. We omit j from $g_i^{(j)}$ and ρ_j to simplify the notation. Let $\tilde{b}_{i,p}$ be the estimate of g_i after decoding it for p levels where $p \in [m]$. Therefore, the estimator $\tilde{g} = \sum_{i=1}^m \frac{\tilde{b}_{i,\rho_i} B}{2^{\rho_i - 1}}$. Let $\tilde{g}_i = \sum_{k=1}^m \frac{\tilde{b}_{i,k} B}{2^{k-1}}$ be the decoded value of g_i till level m and $\bar{g} = \frac{1}{m}\sum_{i=1}^m \tilde{g}_i = \sum_{k=1}^m \frac{\bar{b}_k B}{2^{k-1}}$, where $\bar{b}_k = \frac{1}{m}\sum_{i=1}^m \tilde{b}_{i,k}$.

We compute the expected error for coordinate j, where the expectation is wrt the permutation ρ_j. Note that $\mathbb{E}_{\rho\sim\Pi_m}[\tilde{g}_i] = \bar{g}$.

$$
\begin{aligned}
\mathbb{E}_{\rho\sim\Pi_m}[|g - \tilde{g}|] &= \sqrt{(\mathbb{E}_{\rho\sim\Pi_m}[|g - \tilde{g}|])^2} \leq \sqrt{\mathbb{E}_{\rho\sim\Pi_m}|g - \tilde{g}|^2} \leq \sqrt{\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g} - \bar{g}|^2 + |g - \bar{g}|^2} \\
&\leq \sqrt{\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g} - \bar{g}|^2} + |g - \bar{g}| \leq \frac{1}{m}\sum_{i=1}^m |g_i - \tilde{g}_i| + \sqrt{\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g} - \bar{g}|^2} \\
&\leq \frac{B}{2^{m-1}} + \sqrt{\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g} - \bar{g}|^2}
\end{aligned}
$$

We use Jensen's inequality for the first inequality. For the second inequality, we use bias-variance decomposition for the random variable $\tilde{g}$, where the first term is its variance, and the second term is its bias wrt the term g. We then use $\sqrt{a+b} \leq \sqrt{a} + \sqrt{b}$ for any $a, b \geq 0$. To handle the term $|g - \bar{g}|$, we expand both terms as a summation over m clients, followed by a triangle inequality. As each estimator $\tilde{g}_i$ is at least $\frac{B}{2^{m-1}}$ away from g_i, each term in the difference $|g_i - \tilde{g}_i|$ has the upperbound $\frac{B}{2^{m-1}}$.

We now bound the variance term separately. Note that

$$
\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g} - \bar{g}|^2 = \mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}|^2 - \bar{g}^2
$$

We first evaluate the second moment $\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}|^2$.

$$\begin{aligned}
\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}|^2 &= \mathbb{E}_{\rho\sim\Pi_m}\left|\sum_{i=1}^{m}\frac{\tilde{b}_{i,\rho_i}}{2^{\rho_i-1}}\right|^2 = \sum_{i=1}^{m}\mathbb{E}_{\rho\sim\Pi_m}\left[\frac{\tilde{b}_{i,\rho_i}^2]B^2}{2^{2\rho_i-2}}\right] + B^2\sum\sum_{1\le i\ne j\le m}\mathbb{E}_{\rho\sim\Pi_m}\left[\frac{\tilde{b}_{i,\rho_i}}{2^{\rho_i-1}}\frac{\tilde{b}_{j,\rho_j}}{2^{\rho_j-1}}\right] \\
&= \sum_{k=1}^{m}\frac{B^2}{2^{2k-2}} + B^2\sum\sum_{1\le i\ne j\le m}\mathbb{E}_{\rho_i}\left[\mathbb{E}_{\rho\sim\Pi_m}\left[\frac{\tilde{b}_{i,\rho_i}}{2^{\rho_i-1}}\frac{\tilde{b}_{l,\rho_j}}{2^{\rho_j-1}}|\rho_i\right]\right] \\
&= \sum_{k=1}^{m}\frac{B^2}{2^{2k-2}} + B^2\sum\sum_{1\le i\ne j\le m}\mathbb{E}_{\rho_i}\left[\frac{\tilde{b}_{i,\rho_i}}{2^{\rho_i-1}}\frac{1}{m-1}\sum_{l=1,l\ne\rho_i}^{m}\frac{\tilde{b}_{j,l}}{2^{l-1}}\right] \\
&= \sum_{k=1}^{m}\frac{B^2}{2^{2k-2}} + \frac{B^2}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\left[\frac{\tilde{b}_{i,k}}{2^{k-1}}\sum_{l=1,l\ne k}^{m}\frac{\tilde{b}_{j,l}}{2^{l-1}}\right] \\
&= \sum_{k=1}^{m}\frac{B^2}{2^{2k-2}} + \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\left(\sum_{k=1}^{m}\frac{\tilde{b}_{i,k}B}{2^{k-1}}\right)\left(\sum_{l=1}^{m}\frac{\tilde{b}_{j,l}B}{2^{l-1}}\right) \\
&\quad - \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\frac{B^2\tilde{b}_{i,k}\tilde{b}_{j,k}}{2^{2k-2}} \\
&= \sum_{k=1}^{m}\frac{B^2}{2^{2k-2}} + \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\tilde{g}_i\tilde{g}_j - \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\frac{B^2\tilde{b}_{i,k}\tilde{b}_{j,k}}{2^{2k-2}} \\
&= \frac{m^2|\bar{g}|^2-\sum_{i=1}^{m}|\tilde{g}_i|^2}{m(m-1)} + \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\frac{B^2(|\tilde{b}_{i,k}|^2+|\tilde{b_{j,k}}|^2-2\tilde{b}_{i,k}\tilde{b}_{j,k})}{2^{2k-1}} \\
&= \frac{m}{m-1}|\bar{g}|^2 - \frac{\sum_{i=1}^{m}|\tilde{g}_i|^2}{m(m-1)} + \frac{1}{2m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\left(\frac{B(\tilde{b}_{i,k}-\tilde{b}_{j,k})}{2^{k-1}}\right)^2
\end{aligned}$$

Note that we expand the square of the sum of terms where $\tilde{b}_{i,j}^2 = 1$. For the second term, we use the law of total expectation by conditioning on the value of ρ_i. To evaluate the inner expectation, we note that ρ_j can take any value other than that of ρ_i with equal probability. To evaluate the outer expectation, note that ρ_i can take any value in $[m]$ with equal probability. In the fourth equation, we subtract the term where $l = k$. Then, we can factorize the remaining terms to obtain $\tilde{g}_i$ and $\tilde{g}_j$. Note that the sum of the product terms $\tilde{g}_i\tilde{g}_j$ can be expressed as $|\sum_{i=1}^{m}\tilde{g}_i|^2$, with the square terms subtracted. Further, we express the term $\frac{B^2}{2^{2k-2}} = \sum\sum_{1\le i\ne j\le m}\frac{B^2(|\tilde{b}_{i,k}|^2+|\tilde{b}_{j,k}|^2}{2^{2k-1}}$ as $|\tilde{b}_{i,k}|^2 = 1$. Finally, we complete the squares for each term k.

Using the above value of second moment $\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}|^2$, we can compute the variance,

$$\begin{aligned}
\mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}-\bar{g}|^2 = \mathbb{E}_{\rho\sim\Pi_m}|\tilde{g}|^2 - |\bar{g}|^2 &= \frac{|\bar{g}|^2 - \frac{1}{m}\sum_{i=1}^{m}|\tilde{g}_i|^2}{m-1} + \frac{1}{2m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\left(\frac{B(\tilde{b}_{i,k}-\tilde{b}_{j,k})}{2^{k-1}}\right)^2 \\
&= \frac{1}{2m^2}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\left(\frac{B(\tilde{b}_{i,k}-\tilde{b}_{j,k})}{2^{k-1}}\right)^2
\end{aligned}$$

We use $\bar{g}^2 \le \frac{1}{m}\sum_{i=1}^{m}|\tilde{g}_i|^2 = \frac{1}{2m^2}\sum\sum_{1\le i\ne j\le m}(\tilde{g}_i-\tilde{g}_j)^2 \ge \frac{1}{2m^2}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^{m}\left(\frac{B(\tilde{b}_{i,k}-\tilde{b}_{j,k})}{2^{k-1}}\right)^2$.

To simplify this bound, we need to incorporate the difference in the actual gradient vectors. For this purpose, we try to bound the differences $|\tilde{b}_{i,k}-\tilde{b}_{j,k}|$ in terms of $\Delta_{ij} \triangleq |g_i - g_i|$. If $\Delta_{ij} = |g_i - g_j|$, then $\tilde{b}_{i,k} = \tilde{b}_{j,k}, \forall k \ge \log\left(\frac{B}{\Delta_{ij}}\right)$.

B.2 Proof for Theorem 3

We first show how to obtain the coefficients c_i and the bound on the term Λ in Eq (7). To obtain the coefficients c_i, we replace set $L = m, n = d, R = \frac{m \log L}{d}$ and $\sigma^2 = \frac{B^2}{d}$ in (Venkataramanan et al., 2014b, Eq 2). Note that the terms δ_1 and δ_2 can be directly applied in (Venkataramanan et al., 2014b, Section V) with the same values of ϵ_k and γ_k. Note that the bound on Λ in (Venkataramanan et al., 2014b, Theorem 1) is for gaussian sources with a fixed variance, however, their proof in (Venkataramanan et al., 2014b, Section V) only uses the fact that the squared ℓ_2 norm of the source is bounded, which is exactly our case from Assumption 2.

The proof of Eq (7) is same as Theorem 2 for a single dimension, with the coefficients $\frac{B}{2^{j-1}}$ replaced by c_j and $\tilde{b}^{(r)}_{i,k}$ replaced by $A_{(k-1)L+\tilde{b}_{i,k}}$. Following Appendix B.1, we can write down the ℓ_2 error.

$$\mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g} - g||_2^2] = \mathbb{E}_{\rho\sim\Pi_m}[||g - \mathbb{E}_{\rho\sim\Pi_m}[\tilde{g}]||_2^2] + \mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g} - \mathbb{E}_{\rho\sim\Pi_m}[\tilde{g}]||_2^2]$$

$\mathbb{E}[\tilde{g}] = \bar{g} = \frac{1}{m}\sum_{i=1}^m \bar{g}_i$, where $\bar{g}_i = \sum_{j=1}^m c_j A_{(j-1)L+\tilde{b}_{i,j}}$. By triangle inequality, the first term is $\frac{1}{m}\sum_{i=1}^m ||g_i - \bar{g}_i||_2^2$, which is bounded individually by $\Lambda = B^2(1 + \frac{10\log L}{d} e^{\frac{m\log L}{d}}(\delta_1 + \delta_2))^2 \left(1 - \frac{2\log L}{d}\right)^m$ by setting $L = m, n = d, R = \frac{m\log L}{d}, \sigma^2 = \frac{B^2}{d}$ and $\delta_0 = 0$ in (Venkataramanan et al., 2014b, Theorem 1).

For the second term, we need to bound $\mathbb{E}[||\tilde{g}||_2^2]$.

$$\begin{aligned}
\mathbb{E}[||\tilde{g}||_2^2] &= \frac{1}{m}\sum_{i=1}^m\sum_{j=1}^m c_i^2 ||A_{(j-1)L+\tilde{b}_{i,j}}||_2^2 \\
&\quad + \sum\sum_{1\le i\ne j\le m} \mathbb{E}_{\rho\sim\Pi_m}\left[c_{\pi(i)}c_{\pi(j)}\langle A_{(\pi(i)-1)L+\tilde{b}_{i,\pi(i)}}, A_{(\pi(j)-1)L+\tilde{b}_{j,\pi(j)}}\rangle\right] \\
&= \frac{1}{m}\sum_{i=1}^m\sum_{j=1}^m c_i^2 ||A_{(j-1)L+\tilde{b}_{i,j}}||_2^2 \\
&\quad + \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m} \mathbb{E}_{\rho\sim\Pi_m}\left[c_{\pi(i)}c_{\pi(j)}\langle A_{(\pi(i)-1)L+\tilde{b}_{i,\pi(i)}}, A_{(\pi(j)-1)L+\tilde{b}_{j,\pi(j)}}\rangle\right] \\
&= \frac{m^2||\bar{g}||_2^2 - \sum_{i=1}^m ||\tilde{g}_i||_2^2}{m(m-1)} + \frac{1}{m(m-1)}\sum\sum_{1\le i\ne j\le m}\sum_{k=1}^m c_k^2 ||A_{(k-1)L+\tilde{b}_{j,k}} - A_{(k-1)L+\tilde{b}_{i,k}}||_2^2
\end{aligned}$$

The remainder of the proof follows proof of Theorem 2 with $|\cdot|^2$ replaced by $||\cdot||_2^2$.

For the first equation in Eq (7), since the collaborative compression scheme has precision Λ, for each client, the vector g_i lies in a ball around g of radius $\Delta_{2,\max} = \max_{i\in[m]}||g_i - g||_2$. Therefore, if we use sparse regression codes to individually encode each vector g_i, it's decoded value $\tilde{g}_i$ must lie either in the same ball or be at most Λ ℓ_2 distance away from some vector in this ball. Since this set is convex, and the vector $\bar{g}$ is also obtained as a convex combination of vectors inside this ball or at most Λ away from it, the vector $\bar{g}$ also lies inside this ball or is at most Λ away from it.

This proves Eq (7).

B.3 Proof of Equation (8) and Remark 1

To bound Δ_{reg}, we need to bound the terms of the form $||A_{(k-1)L+\tilde{b}_{i,k}} - A_{(k-1)L+\tilde{b}_{j,k}}||_2^2$ for clients $i \ne j \in [m]$ and at level $k \in [m]$. Note that if $\tilde{b}_{j,k} = \tilde{b}_{i,k}$, then this difference is 0. If $\tilde{b}_{j,k} \ne \tilde{b}_{i,k}$, then this is Γ^2 from the

definition of $(\delta_1, \delta_2, \Gamma)$-Cover. This results in the following upper bound for Δ_{reg}.

$$\Delta_{reg} \leq \frac{\Gamma^2}{m} \sum_{i,j\in[m], i\neq j} \sum_{k=1}^{m} c_k^2 \cdot \mathbf{1}(\tilde{b}_{i,k} \neq \tilde{b}_{j,k}) = \frac{2\Gamma^2 B^2 \log L}{d^2 m} \sum_{i,j\in[m], i\neq j} \sum_{k=1}^{m} \left(1 - \frac{2\log L}{d}\right)^{k-1} \mathbf{1}(\tilde{b}_{i,k} \neq \tilde{b}_{j,k})$$
$$= \frac{2\Gamma^2 B^2 \log L}{dm^2} \sum_{i,j\in[m], i\neq j} \sum_{k=1}^{m} \left(1 - \frac{2\log L}{d}\right)^{k-1} \mathbf{1}(\tilde{b}_{i,k} \neq \tilde{b}_{j,k})$$

We use the value of c_k in the second inequality and then simplify the terms.

To prove the bound of Γ for the case of Gaussian A in Remark 1, note that Γ is an upper bound on the the squared ℓ_2 norm of the difference of two gaussian random vectors. For any $r_1, r_2 \in [L]$ with $r_1 \neq r_2$, $A_{(k-1)L+r_1} - A_{(k-1)L+r_2} \sim \mathcal{N}(0, 2\mathbb{I}_d)$ as each vector is d-dimensional, they are sampled independently, and each element in each vector is sampled from a unit normal. Therefore, $\frac{1}{2}||A_{(k-1)L+r_1} - A_{(k-1)L+r_2}||_2^2$ is a χ_d^2 random variable. Note that the scaling of $\frac{1}{2}$ is to normalize the covariance to $\mathbb{I}_d$. By concentration of χ_d^2 random variables from (Laurent & Massart, 2000), we know that, with probability $1-\delta$,

$$||A_{(k-1)L+r_1} - A_{(k-1)L+r_2}||_2^2 \leq 2d + 4\sqrt{d\log(1/\delta)} + 4\log(1/\delta)$$

Since $r_1, r_2 \in [L]$, we can take a union bound over all $\frac{L(L-1)}{2}$ unique pairs of (r_1, r_2). Further, we can take a union bound over all levels $k \in [m]$. These would set the probability of error to $\frac{mL(L-1)\delta}{2}$. Setting δ_3 to this value, we obtain $\delta = \frac{2\delta_3}{mL(L-1)}$ and the following bound $\forall k \in [m], r_1 \neq r_2 \in [L]$,

$$||A_{(k-1)L+r_1} - A_{(k-1)L+r_2}||_2^2 \leq 2d + 4\sqrt{d\log(\frac{mL^2}{\delta_3})} + 4\log(\frac{mL^2}{\delta_3}) \triangleq \Gamma^2$$

We upper bound $\frac{L(L-1)}{2}$ by L^2.

B.4 Improving HadamardMultiDim and SparseReg with Repetitions

Let $\tilde{g}_R = \frac{1}{R}\sum_{r=1}^{R} \tilde{g}(r)$ be estimator after R repetitions where $\tilde{g}(r)$ is the estimator after 1 repetition of either HadamardMultiDim or SparseReg. To compute the error for both Theorem 2 and Theorem 3, we use a bias-variance decomposition of the squared estimation error. Consider SparseReg where we bound $\mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g} - g||_2^2]$. Note that even with the new definition of $\tilde{g}$, it's mean wrt ρ remains the same as $\mathbb{E}_{\rho\sim\Pi_m}[\tilde{g}] = \frac{1}{R}\sum_{r=1}^{R} \mathbb{E}_{\rho\sim\Pi_m}[\tilde{g}(r)] = \frac{1}{R}\sum_{r=1}^{R} \bar{g} = \bar{g}$. This contributes to the term with exponential dependence on m and it remains the same. Note that the variance term, however, decreases with R.

$$\mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g} - \bar{g}||_2^2] = \frac{1}{R^2}\sum_{r=1}^{R} \mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g}(r) - \bar{g}||_2^2] = \frac{1}{R}\mathbb{E}_{\rho\sim\Pi_m}[||\tilde{g}(1) - \bar{g}||_2^2] = \frac{\Delta_{\text{reg}}}{R}$$

Similarly for HadamardMultiDim, the term Δ_{Hadamard} is replaced by $\frac{\Delta_{\text{Hadamard}}}{\sqrt{R}}$, as we take square root of the variance in it's analysis. Note that the $\Delta_{2,\max}$ and $\Delta_{\infty,\max}$ terms are not affected by R as it's analysis does not use ρ.

C Proofs for Section 4

C.1 Proof of Lemma 1

To prove this Lemma, note that $\tilde{b}_i = sign(\langle g_i, z_i\rangle) \neq sign(\langle g, z_i\rangle)$ only if z_i is sampled from the symmetric difference of g_i and g. The probability that a z_i sampled uniformly from $\mathbb{S}^{d-1}$ lies in this symmteric difference is given by $\arccos(\langle g, g_i\rangle)/\pi$. If we set $\Delta_{\text{corr}} = \frac{1}{m\pi}\sum_{i\in[m]} \arccos(\langle g, g_i\rangle)$

Let ζ be the fraction of z_i such that $\tilde{b}_i \neq sign(\langle g, z_i\rangle)$. Then, by Chernoff bound, we have,

$$\Pr[\zeta \geq (1+\gamma)\Delta_{\text{corr}}] \leq e^{-\frac{\gamma^2 m \Delta_{\text{corr}}}{2+\gamma}}$$

By setting γ to be any small constant, we obtain, with probability $1-\mathcal{O}(\exp(-m\Delta_{\text{corr}}))$, atmost $\zeta = \Theta(\Delta_{\text{corr}})$ fraction of datapoints are not generated from the halfspace with normal g and are thus corrupted.

C.2 Proofs of Theorem 4 and 5

To prove Theorem 4, we utilize the guarantees of (Shen, 2023, Theorem 3), where the sample complexity requirement ensures that the error is $\tilde{O}(\frac{d}{m})$. Further, (Shen, 2023, Theorem 3) obtains error guarantee linear in the noise rate of the samples which is obtained from Lemma 1. The error guarantee is in terms of the symmetric difference between $\tilde{g}$ and g wrt the uniform distribution on the unit sphere. Since this is equal to the angle between these two vectors divided by π, this gives us a bound on the inner product of these two unit vectors.

To prove Theorem 5, from (Kalai et al., 2008, Theorem 12), the sample complexity provides the term $\frac{d}{\sqrt{m}}$ while the noise tolerance provides the term $\sqrt{d}\Delta_{\text{corr}}$.

C.3 Proof of Equation (11)

To prove this remark, note that $\arccos(x)$ is concave for $x \geq 0$. Therefore, by applying Jensen's inequality, we obtain,

$$\begin{aligned}
\Delta_{\text{corr}} =& \frac{1}{m\pi}\sum_{i\in[m]} \arccos(\langle g_i, g\rangle) \leq \frac{1}{\pi}\arccos\left(\langle \frac{1}{m}\sum_{i=1}^{m} g_i, g\rangle\right) = \frac{1}{\pi}\arccos\left(||\frac{1}{m}\sum_{i=1}^{m} g_i||_2 \langle g, g\rangle\right) \\
\leq& \frac{1}{\pi}\arccos\left(\sqrt{||\frac{1}{m}\sum_{i\in[m]} g_i||_2^2}\right) = \frac{1}{\pi}\arccos\left(\sqrt{||\frac{\sum_{i\in[m]}\langle g_i, g_i\rangle}{m^2} + \frac{2}{m^2}\sum\sum_{1\leq i<j\leq m}\langle g_i, g_j\rangle||}\right) \\
=& \frac{1}{\pi}\arccos\left(\sqrt{\frac{1}{m} + \frac{2}{m^2}\sum\sum_{1\leq i<j\leq m}\langle g_i, g_j\rangle}\right)
\end{aligned}$$

D Additional Experiment Details

Baselines We implement all the baselines mentioned in Table 2. As all these baselines are suited to ℓ_2 error, for the DME experiment on gaussians, where ℓ_2 error is the correct metric, compare SparseReg (Algorithm 4) to all these baselines. For ℓ_∞ error uniform distribution, we implement NoisySign (Algorithm 1) and HadamardMultiDim (Algorithm 3) and compare it to Correlated SRQ (Suresh et al., 2022), as it's guarantees hold in single dimensions. We also add comparisons to its independent variant, SRQ (Suresh et al., 2017), and Drive (Vargaftik et al., 2021), which performs coordinate-wise signs. For the unit vector case, we implement OneBit (Algorithm 5 Technique II) and SparseReg(Algorithm 4) and compare it with one independent compressor (SRQ (Suresh et al., 2017)) and one collaborative compressor (RandKSpatialProj (Jiang et al., 2023)). Note that we set $d = 512$ throughout our experiments and tune the parameters (number of coordinates sent (Konečný & Richtárik, 2018; Jhunjhunwala et al., 2021) or the quantization levels in (Suresh et al., 2017; 2022)) so that all compressors have the same number of bits communicated. For compressors without tunable parameters, we repeat them to match the communication budget. The communication budget is 2375 ± 25 bits/client for each compressor every round. It is often not possible for two different forms of compressors, like sparsification methods and quantization methods to achieve the exact same communication budget even after tuning hence we allow a small range of communication budgets.

Datasets For the distributed mean estimation task, we generate d dimensional vectors on $m = 100$ clients. To compare ℓ_2 error, we generate g with $||g||_2 = 100$. Then, each client generates g_i from a $\mathcal{N}(0, \Delta_2^2)$, where $\Delta_2 \in [0.001, 100]$. To compare ℓ_∞ error, we generate g uniformly from a hypercube $[-B, B]^d$ where $B = 100$. Each client generates g_i from a smaller hypercube $[-\Delta_\infty, \Delta_\infty]^d$ centered at g where $\Delta_\infty \in [10^{-3}, 10^2]$. To compare cosine distance, we generate g uniformly from the unit sphere, and each client generates g_i uniformly from the set of unit vectors at a cosine distance Δ_{corr} from the g, Here, $\Delta_{corr} \in [0.01, 0.4]$.

For KMeans and power iteration, we set $m = 50$. FEMNIST is a real federated dataset where each client has handwritten digits from a different person. We apply dimensionality reduction to set $d = 512$. We run 20 iterations of Lloyd's algorithm (Lloyd, 1982) for KMeans and 30 power iterations. For distributed linear regression, the Synthetic dataset is a mixture of linear regressions, with one mixture component per client. The true model $w_i \in \mathbb{R}^d$ for each component is obtained from DME setup for gaussians with $\Delta_2 = 4$. Then, we generate $n = 1000$ datapoints on each client, where the features x are sampled from standard normal, while the labels y are generated as $y = \langle w_i, x\rangle + \xi$, where ξ is the zero-mean gaussian noise with variance 10^{-2}. For UJIndoorLoc, we use the first $d = 512$ of the 520 features following (Jiang et al., 2023). The task for UJIndoorLoc dataset is to predict the longitude of a phone call. For both the linear regression datasets, we run 50 iterations of GD. For MNIST and UJIndoorLoc, we split the dataset uniformly into m chunks one per client.

For logistic regression for binary classification, we select the last 2 classes of the HAR and label them with ± 1. We split the dataset into $m = 20$ clients iid. HAR dataset has 561 features which we reduce by PCA to $d = 512$. We run distributed gradient descent with learning rate 0.001 for $T = 200$ iterations on the logistic loss, where the logistic loss for any data point $(x, y) \in \mathbb{R}^d \times \{\pm 1\}$ is defined as $\ell(w, (x, y)) = \log(1 + \exp(-\langle w, x\rangle \cdot y)$ for any weight $w \in \mathbb{R}^d$.

Metrics With the same number of bits, we can directly compare the error of baselines. For mean estimation, we measure ℓ_2 error, ℓ_∞ error and cosine distance for gaussian, uniform and unit vectors respectively. For KMeans, we report the KMeans objective. For power iteration, we report the top eigenvalue. For linear regression, we provide the mean squared error on a test dataset. For logistic regression, we report the training logistic loss and test accuracy for binary classification. For all experiments except power iteration and test accuracy logistic regression, lower value of the reported metric implies better performance. For power iteration, higher implies better performance, as we need to find the eigenvector corresponding to the top eigenvalue. For logistic regression, higher test accuracy implies better performance.

We provide the code in the supplementary material and all the experiments took 6 days to run on a single 20 core machine with 25 GB RAM.

E Distributed Gradient Descent with SparseReg Compressor

This section uses our ℓ_2 compressor, SparseReg, for running FedAvg. Each client $i \in [m]$ contains a local objective function $f_i : \mathcal{W} \to \mathbb{R}$. We define the global objective function $f(w) = \frac{1}{m}\sum_{i=1}^{m} f_i(w)$, $\forall w \in \mathcal{W} \subset \mathbb{R}^d$. The goal is to find $w^\star \in \arg\min_{w\in\mathcal{W}} f(w)$. Note that $\nabla f(w) = \frac{1}{m}\sum_{i=1}^{m} \nabla f_i(w)$, therefore, in our case, the vector g_i correspond to $\nabla f_i(w)$. We describe the algorithm in Algorithm 6

Algorithm 6 Distributed Projected Gradient Descent with SparseReg compressor

Require: Initial iterate $w^0 \in \mathcal{W}$, Step size $\gamma > 0$
  Server
  SparseReg-`Init()`
  for $t = 0$ to $T - 1$ **do**
    Send w^t to all clients $i \in [m]$.
    Receive $\tilde{b_i}^t$ from clients $i \in [m]$.
    $\tilde{g}^t \leftarrow$ SparseReg-`Decode`$(\{\tilde{b}_i^t\}_{i\in[m]})$
    $w^{t+1} \leftarrow \text{proj}_{\mathcal{W}}(w^t - \eta_t \tilde{g}^t)$
  end for
  `Client(i)` at iteration t
  Receive w^t from server.
  $\tilde{b}_i \leftarrow$ SparseReg-`Encode`$(\nabla f_i(w^t))$
  Send $\tilde{b}_i^t$ to server.

We first state the assumptions required for applying the SparseReg compressor.

Assumption 4 (Bounded Gradient)**.** *For all $w \in \mathcal{W}, i \in [m]$, we assume that $||\nabla f_i(w)||_2 \leq B$.*

By this assumption, we ensure that for each iteration t in Algorithm 6, $||g_i||_2 = ||\nabla f_i(w^t)||_2$ is bounded. Further, bounded gradients imply that each f_i is Lipschitz. By triangle inequality, we can also establish the following corollary.

Corollary 1. *The objective function $f(w)$ is B-Lipschitz, $\forall w \in \mathcal{W}$.*

From the above assumptions, it is clear that local objective functions need to be Lipschitz. From (Bubeck, 2015, Theorem 3.2), if the domain of iterates, $\mathcal{W}$ is bounded and $f(w)$ is also convex, then gradient descent can converge at a rate $\mathcal{O}(1/\sqrt{T})$. We use these two assumptions, and establish a $\mathcal{O}(1/\sqrt{T})$ rate along with a error obtained from Theorem 3. We define $\Delta_{\text{reg}}(t)$ and $\Delta_{2,\max}(t)$ from Theorem 3 to be the corresponding errors for $g_i = \nabla f_i(w^t), \forall i \in [m]$ for any $t > 0$.

Assumption 5 (Bounded domain)**.** *The set $\mathcal{W}$ is closed and convex with diameter R^2.*

Assumption 6 (Convexity)**.** *The objective function $f(w)$ is convex $\forall w \in \mathcal{W}$.*

We now state our convergence result.

Theorem 6. *Under Assumptions 4, 5, 6, running Algorithm 6 for T iterations with step size $\eta_t = \frac{R}{B\sqrt{T}}$, with probability $1 - 2m^2 LT \exp(-d\delta_1^2/8) - mT\left(\frac{L^{2\delta_2}}{\log L}\right)^{-m}$ we have,*

$$\begin{aligned}
&\mathbb{E}[f(\bar{w}^T)] - f(w^\star) \leq \frac{R(2B^2 + \Gamma_1)}{2B\sqrt{T}} + \sqrt{\Gamma_1}R, \quad \textit{where,} \quad \bar{w}^T = \frac{1}{T}\sum_{t=0}^{T-1} w^t \\
&\Gamma_1 = B^2\left(1 + \frac{10\log L}{d}\exp\left(\frac{m\log L}{d}\right)(\delta_1 + \delta_2)\right)^2\left(1 - \frac{2\log L}{d}\right)^m, \\
&\Gamma_2 = \max_{t\in\{0,1,\ldots,T-1\}} \min\{\Delta_{\text{reg}}(t), \Delta_{2,\max}(t)\}
\end{aligned} \tag{13}$$

From the above theorem, we can see that the high probability terms and Γ_1 and Γ_2 are obtained from Theorem 3. Note that $\Gamma = \mathcal{O}(B^2\exp(-m/d))$, therefore, for large m, the additional bias term of $R\sqrt{\Gamma_1}$ is very small. Further, the term $\Gamma_2 \leq B^2$, therefore, Γ_2 only affects constant terms in the convergence rate due to $\sqrt{T}$ in the denominator. If $\exp(-m/d) = \mathcal{O}(1/\sqrt{T})$ or $m = \Omega(d\log T)$, the final convergence rate of Algorithm 6 is $\mathcal{O}(RB/\sqrt{T})$ which is the rate for distributed GD without compression.

We provide the proof for the above theorem, which modifies the proof of (Bubeck, 2015, Theorem 3.2) to handle a biased gradient oracle. We can also extend our analysis to other function classes, for instance strongly convex functions, by using existing works on biased gradient oracles (Ajalloeian & Stich, 2020). Extending the proof to FedAvg from distributed GD would require using biased gradient oracles in (Li et al., 2020). Further, these proofs can also be extended to HadamardMultiDim compressor, with an additional $\sqrt{d}$ factor in the corresponding error terms from Theorem 2 to account for conversion from ℓ_∞ to ℓ_2 norm.

E.1 Proof of Theorem 6

At any iteration $t > 0$, we use $\tilde{g}^t$ to denote the estimate of $\nabla f(w^t)$. From the proof of Theorem 3, $||\mathbb{E}_t[\tilde{g}^t] - \nabla f(w^t)||_2 \leq \sqrt{\Gamma_1}$, and $\mathbb{V}ar_t(\tilde{g}^t|w^t) \leq \Gamma_2, \forall t > 0$, where $\mathbb{E}_t$ and $\mathbb{V}ar_t$ are the expectation and variance wrt the randomness in the SparseReg compressor at iteration t. We take a union bound over the high probability terms in Theorem 3 over all iterations $t = 0$ to $T - 1$.

We can write the following equation by convexity of $f(w^t)$.

$$\begin{aligned}
f(w^t) - f(w^\star) \leq& \langle \nabla f(w^t), w^t - w^\star\rangle = \langle \tilde{g}^t, w^t - w^\star\rangle + \langle \nabla f(w^t) - \tilde{g}^t, w^t - w^\star\rangle \\
\leq& \frac{1}{2\eta}(||w^t - w^\star||_2^2 - ||w^t - \eta\tilde{g}^t - w^\star||_2^2) + \eta||\tilde{g}^t||_2^2/2 + \langle \nabla f(w^t) - \tilde{g}^t, w^t - w^\star\rangle
\end{aligned}$$

In the second line, we use $2\langle a, b\rangle = ||a||_2^2 + ||b||_2^2 - ||a-b||_2^2$. Now, taking expectation wrt the randomness in SparseReg at iteration t, we obtain,

$$\begin{aligned}
\mathbb{E}_t[f(w^t)] - f(w^\star) \leq & \frac{1}{2\eta}(||w^t - w^\star||_2^2 - \mathbb{E}_t[||w^t - \eta\tilde{g}^t - w^\star||_2^2]) + \eta\mathbb{E}_t[||\tilde{g}^t||_2^2]/2 \\
& + \langle \nabla f(w^t) - \mathbb{E}_t[\tilde{g}^t], w^t - w^\star\rangle \\
\leq & \frac{1}{2\eta}(||w^t - w^\star||_2^2 - \mathbb{E}_t[||w^{t+1} - w^\star||_2^2]) + \eta(||\mathbb{E}_t[\tilde{g}^t]||_2^2 + \mathbb{V}ar_t(\tilde{g}^t))/2 \\
& + ||\nabla f(w^t) - \mathbb{E}_t[\tilde{g}^t]||_2 \cdot ||w^t - w^\star||_2 \\
\leq & \frac{1}{2\eta}(||w^t - w^\star||_2^2 - \mathbb{E}_t[||w^{t+1} - w^\star||_2^2]) + \eta(B^2 + \Gamma_2)/2 + \sqrt{\Gamma_1}R
\end{aligned}$$

In the second line, we use the non-expansiveness of projections on a convex set, $||w^t - \eta\tilde{g}^t - w^\star||_2 \geq ||\text{proj}_{\mathcal{W}}(w^t - \eta\tilde{g}^t - w^\star)||_2$, the decomposition of 2nd moment into square of mean and variance, and cauchy-schwartz inequality. In the third line, we plug in bounds of Γ_1, Γ_2, diameter of the set and by triangle inequality, argue that $\mathbb{E}[\tilde{g}^t]$ also lies in an ℓ_2 ball of radius B.

Finally, we take expectations wrt all random variables, unroll the recursion from $t = 0$ to T, and divide both sides by T.

$$\frac{1}{T}\sum_{t=0}^{T} \mathbb{E}[f(w^t)] - f(w^\star) \leq \frac{R^2}{2\eta T} + \frac{\eta(B^2 + \Gamma_2)}{2} + \sqrt{\Gamma_1}R \leq \frac{R(2B^2 + \Gamma_1)}{2B\sqrt{T}} + \sqrt{\Gamma_1}R$$

We obtain the final inequality by plugging in the step size $\eta = \frac{R}{B\sqrt{T}}$. By convexity of f, for $\bar{w}^T = \sum_{t=0}^{T-1} w^t$, we obtain,

$$\mathbb{E}[f(\bar{w}^T)] - f(w^\star) \leq \frac{1}{T}\sum_{t=0}^{T-1} \mathbb{E}[f(w^t)] - f(w^\star) \leq \frac{R(2B^2 + \Gamma_1)}{2B\sqrt{T}} + \sqrt{\Gamma_1}R$$