GRANGER CAUSAL INFERENCE ON DAGS IDENTIFIES GENOMIC LOCI REGULATING TRANSCRIPTION

ABSTRACT

When a dynamical system can be modeled as a *sequence* of observations, Granger causality is a powerful approach for detecting predictive interactions between its variables. However, traditional Granger causal inference has limited utility in domains where the dynamics need to be represented as directed acyclic graphs (DAGs) rather than as a linear sequence, such as with cell differentiation trajectories. Here, we present GrID-Net, a framework based on graph neural networks with lagged message passing for Granger causal inference on DAG-structured systems. Our motivating application is the analysis of single-cell multimodal data to identify genomic loci that mediate the regulation of specific genes. To our knowledge, GrID-Net is the first single-cell analysis tool that accounts for the temporal lag between a genomic locus becoming accessible and its downstream effect on a target gene's expression. We applied GrID-Net on multimodal single-cell assays that profile chromatin accessibility (ATAC-seq) and gene expression (RNA-seq) in the same cell and show that it dramatically outperforms existing methods for inferring regulatory locus–gene links, achieving up to 71% greater agreement with independent population genetics-based estimates. The loci identified by GrID-Net are also enriched for 10–50x more transcription-factor binding motifs than other methods. By extending Granger causality to DAG-structured dynamical systems, our work unlocks new domains for causal analyses and, more specifically, opens a path towards elucidating gene regulatory interactions relevant to cellular differentiation and complex human diseases at unprecedented scale and resolution.

1 INTRODUCTION

Understanding the structure of a multivariate dynamical system often boils down to deciphering the causal relationships between its variables. Since inferring true causality is often infeasible, requiring additional experiments or independent mechanistic insight, statistical analysis of observational data to identify predictive relationships between the system's variables can be very valuable. The framework of Granger causality does exactly that: in a dataset where observations are temporally ordered, a time-dependent variable x (with value x_t at time t) is said to "Granger cause" a second variable y if the history of x at time t (i.e., $x_1, \ldots, x_{t-1}$) is significantly predictive of y at time t even after accounting for y's own history (Granger, 1969; Shojaie & Fox, 2021). Originating in econometrics, Granger causality has been a powerful tool in many domains including biology, finance, and social sciences (Fujita et al., 2007; Yao et al., 2015; Benhmad, 2012; Rasheed & Tahir, 2012). The prerequisite for applying Granger causality, however, is that there be a clear sequential ordering of observations, i.e., the data must conform to a *total* ordering along time.

Often, only a *partial* ordering of the observations is possible. For instance, cell differentiation trajectories may have branches. In some social network settings, information flows along a tree (e.g., the retweets of a meme by the followers of an earlier retweeter). In text-mining, the citation graph of publications captures the flow of ideas and knowledge. In such cases, the dynamics of the system are more suitably represented as a directed acyclic graph (DAG) corresponding to the partial ordering, with nodes of the DAG representing dynamical states and its edges indicating the flow of information between states.

The key conceptual advance of this work is extending the applicability of Granger causality to systems with partially ordered observations (**Figure** 1a). On a DAG where each node takes on multi-

Figure 1: OVERVIEW (**a**) We extend the concept of Granger causality, previously applied only to sequentially ordered observations, to partial orderings, i.e., where the information flow can be described as a directed acyclic graph (DAG). (**b**) GrID-Net, our approach for inferring Granger causality on a DAG-structured system, is a graph neural network framework with lagged message-passing (L is defined in Sec. 2.2). (**c**) To identify genomic loci whose accessibility Granger causes the expression of a specific gene, we apply GrID-Net on multimodal single-cell data where chromatin accessibility (ATAC-seq) and gene expression (RNA-seq) have been profiled in the same cell. We model the lag between them by relating the RNA-seq readouts of a cell not to the ATAC-seq readouts of the same cell but to ATAC-seq readouts of cells slightly earlier in the profiled biological process, applying GrID-Net on a DAG constructed from the kNN graph of cells with edges oriented as per pseudotime. (**d**) GrID-Net leverages the biological intuition that the accessibility of a specific genomic locus (ATAC-seq) precedes the binding of regulator proteins to the locus, thus causing the expression levels for its targeted gene (RNA-seq) to be changed.

variate values, we define a variable x as Granger causing another variable y if the latter's value y_v at the DAG node v can be significantly predicted by the values of x at the ancestor nodes of v, even after accounting for y_v's own ancestral history. We introduce GrID-Net (**Gr**anger **I**nference on **D**AGs), a graph neural network (GNN) framework for Granger causal inference in dynamical systems that can be represented as DAGs (**Figure** 1b). We modify the standard GNN architecture to enable lagged message-passing on the DAG, enabling us to accumulate past information for a variable according to the DAG's structure. This, combined with a mathematical formulation of Granger causality more amenable to neural networks, allows GrID-Net to recover nonlinear long-range Granger causal dependencies between the variables.

While our problem formulation is fully general and GrID-Net is broadly applicable, our motivating application is inferring non-coding genomic loci that causally influence the expression of a specific gene. We apply GrID-Net to single-cell multimodal studies that assay chromatin accessibility (ATAC-seq) and gene expression (RNA-seq) in the same cell, seeking to identify a Granger causal relationship between the accessibility of a chromatin region (ATAC-seq "peak") and the expression of a specific gene. The system dynamics are represented by a DAG where each node corresponds to a cell and edges exist between cells in similar dynamic states (i.e. similar ATAC-seq and RNA-seq profiles), with the edge directions estimated by a pseudotime analysis (Haghverdi et al., 2016) (**Figure** 1c).

Our work addresses a pressing need to identify causal relationships between non-coding loci and gene expression, as traditional approaches for doing so are expensive and slow. They include population genetics techniques like expression quantitative trait locus (eQTL) studies, in which human genetic variability at a particular genomic locus is associated with expression changes in the gene of interest. Alternatively, perturbation-based approaches directly seek to identify gene expression changes in response to CRISPR alterations of specific non-coding genomic loci (Fulco et al., 2016).

The emergence of single-cell multimodal assays that profile ATAC-seq and RNA-seq simultaneously offers a new way of discovering such regulatory relationships, on which we capitalize. Rather than requiring the large population sizes of eQTL studies or being limited to specific noncoding loci as in the perturbation-based approaches, these multimodal experiments enable unbiased high-throughput, genome-wide estimations of peak–gene associations.

However, leveraging this data will require algorithmic innovations. Over 99.5% of ATAC-seq observations ("peaks") in a cell are reported as empty (Singh et al., 2021; Zhang et al., 2019), and this sparsity constrains current approaches. A key limitation of existing methods is also that they are based on correlation estimates and cannot capture the temporal dynamics governing the causal relationship between chromatin accessibility and gene expression. For a noncoding locus to influence a gene's expression, the locus typically must first become accessible, upon which regulator proteins act on it to drive a change in gene expression (**Figure** 1d). This causal mechanism, therefore, entails temporal asynchrony between the ATAC-seq and RNA-seq modalities (Lara-Astiaso et al., 2014; Ostuni et al., 2013). GrID-Net is specifically designed to recover the dynamics between peaks and genes while being robust to the noise and sparsity in single-cell data.

We apply GrID-Net on three single-cell multimodal datasets that characterize a range of dynamic processes, including cancer drug responses and cellular differentiation. We show that GrID-Net substantially outperforms current methods in identifying known peak–gene links, opening a path towards unlocking critical insights about gene regulation at genome scale. To our knowledge, this work presents both the first framework for Granger causal inference on dynamical systems represented as DAGs and the first single-cell method to explicitly model the temporal asynchrony between chromatin accessibility and gene expression.

2 METHODS

2.1 BACKGROUND ON GRANGER CAUSALITY: STANDARD FORMULATIONS

In time series analyses, the classical Granger causality framework uses the vector autoregressive (VAR) model, which can be expressed as follows (Lütkepohl, 2005).

$$y_t = \sum_{\ell=1}^{L} a_1^{(\ell)} y_{t-\ell} + \sum_{\ell=1}^{L} a_2^{(\ell)} x_{t-\ell} + \epsilon_t \qquad (1)$$

Here, $y_t \in \mathbb{R}$ and $x_t \in \mathbb{R}$ are values at time t for two stationary time series y and x, respectively; $a_1^{(\ell)}$ and $a_2^{(\ell)}$ are coefficients that specify how lag ℓ affects the future of time series y; and ϵ_t is a zero-mean noise term. In this model, time series y is assumed to be a linear combination of the past L lags of y and x each, with x said to Granger cause time series y if and only if $a_2^{(\ell)} \neq 0$ for all ℓ.

A more generalized formulation — the one we follow — for performing Granger causal inference is to consider two related models for forecasting time series y. A *full* model considers the past values of both y and x to forecast y, while a *reduced* model excludes the effect of x, only containing terms related to the past values of y.

$$y_t = f^{(full)}(y_{t-1}, ..., y_L; x_{t-1}, ..., x_{t-L}) + \epsilon_t \qquad (2)$$
$$y_t = f^{(reduced)}(y_{t-1}, ..., y_{t-L}) + \epsilon_t \qquad (3)$$

Here, $f^{(full)}(\cdot)$ and $f^{(reduced)}(\cdot)$ are generalized functions that specify how the value of y_t depends on past values of y and x. The predictions of each of these models are then compared, upon which a Granger causal relationship is declared if the full model's predictions are significantly more accurate than those of the reduced model.

2.2 GRANGER CAUSALITY ON A DAG: GRAPH NEURAL NETWORK FORMULATION

Let the data be represented by a DAG $\mathcal{G} = (\mathcal{V}, \mathcal{E})$ with $n = |\mathcal{V}|$ nodes (i.e., observations) and the directed edges $\mathcal{E}$ indicating the information flow or partial order between these observations. Let $\boldsymbol{y} \in \mathbb{R}^n$ and $\boldsymbol{x} \in \mathbb{R}^n$ correspond to the values of variables y and x on the nodes in $\mathcal{V}$, with x putatively causing y. To infer Granger causality, we compare full and reduced models as above:

$$\boldsymbol{y} = g(\tilde{\boldsymbol{h}}_y^{(full)} + c\tilde{\boldsymbol{h}}_x^{(full)}) + \boldsymbol{\epsilon} \qquad (4)$$

$$\boldsymbol{y} = g(\tilde{\boldsymbol{h}}_y^{(reduced)}) + \boldsymbol{\epsilon} \qquad (5)$$

Here, $\tilde{\boldsymbol{h}}_y^{(full)}, \tilde{\boldsymbol{h}}_y^{(reduced)}, \tilde{\boldsymbol{h}}_x^{(full)} \in \mathbb{R}^n$ represent the historical information of y or x (as denoted by the subscript) that is aggregated by graph neural network (GNN) layers. For example, at each node $v \in \mathcal{V}$, $\boldsymbol{h}[v]$ is the information accumulated from v's ancestors in $\mathcal{G}$. While $\tilde{\boldsymbol{h}}_y^{(full)}$ and $\tilde{\boldsymbol{h}}_y^{(reduced)}$ are outputs determined from similar architectures, their learned weights and biases end up being different since $\tilde{\boldsymbol{h}}_y^{(full)}$ is influenced by its interaction with $\tilde{\boldsymbol{h}}_x$. The coefficient c mediates this interaction and describes the effect of x's history on y. In addition, $\boldsymbol{\epsilon} \in \mathbb{R}^n$ is a zero-mean noise term. We also include $g(\cdot)$ as an optional link function that maps its input to the support of the distribution of variable y. In our single-cell application, we set g as the exponential function since the target variable ($\boldsymbol{y}$) represents normalized transcript counts, which are non-negative.

For brevity, we only describe how $\tilde{\boldsymbol{h}}_x^{(full)}$ is computed, with $\tilde{\boldsymbol{h}}_y^{(full)}$ and $\tilde{\boldsymbol{h}}_y^{(reduced)}$ being computed analogously. Also, since $\boldsymbol{x}$ occurs only in the full model, for notational convenience we drop its superscript *(full)* below, writing just $\tilde{\boldsymbol{h}}_x$. We express $\tilde{\boldsymbol{h}}_x$ as the mean of the outputs of L consecutive GNN layers, denoting the layerwise outputs as $\boldsymbol{h}_x^{(\ell)}$:

$$\tilde{\boldsymbol{h}}_x = \frac{1}{L} \sum_{\ell=1}^{L} \boldsymbol{h}_x^{(\ell)} \qquad (6)$$

$$\boldsymbol{h}_x^{(\ell)} = \begin{cases} f_x^{(\ell)}(\boldsymbol{h}_x^{(\ell-1)}, \boldsymbol{A}_+) & \text{if } \ell > 1 \\ f_x^{(\ell)}(\boldsymbol{x}, \boldsymbol{A}) & \text{if } \ell = 1 \end{cases} \qquad (7)$$

Here, $\boldsymbol{A}$ and $\boldsymbol{A}_+ \in \mathbb{R}^{n \times n}$ are matrices defined by the DAG structure, designed to introduce a one-step lag in the first GNN layer that is propagated by subsequent layers. $\boldsymbol{A}$ is the column-normalized adjacency matrix of $\mathcal{G}$, with $A_{ij} = \frac{1}{d_j}$ if edge $(i, j) \in \mathcal{E}$ and 0 otherwise. d_j is the in-degree of node j, ensuring the column normalization. We note that a DAG does not have self-loops so that the diagonal terms of $\boldsymbol{A}$ are zero. In $\boldsymbol{A}_+$ we introduce these diagonal terms, setting them to the same value as others in the column (d_j is incremented by 1 to preserve normalization).

Together, $\boldsymbol{A}$ and $\boldsymbol{A}_+$ allow us to extend the key intuition of Granger causality to DAGs: predicting a value at node v from the values of its ancestors that are within L steps. The first layer introduces the lag, using information at the parents of v but not the information at v itself. Each subsequent layer introduces the preceding set of parents, with the diagonal term in $\boldsymbol{A}_+$ ensuring that information stored at v from previous layers is integrated with the new information propagated via v's parents. The sequence of GNN layers therefore reflects the successive aggregation of past information by traversing L steps backward along the DAG for each node in the graph. Thus, we are able to match the concept of lags in time series-based formulations (e.g., Eqn. 1) but with the crucial advantage of leveraging the richness of a DAG's structure in directing the information flow.

We compute the GNN layers as follows (again, we only describe f_x for brevity):

$$f_x^{(\ell)}(\boldsymbol{h}_x^{(\ell-1)}, \boldsymbol{A}_+) = \sigma(w_x^{(\ell)} \boldsymbol{A}_+^T \boldsymbol{h}_x^{(\ell-1)} + b_x^{(\ell)}) \quad \text{if } \ell > 1 \qquad (8)$$

$$f_x^{(\ell)}(\boldsymbol{x}, \boldsymbol{A}) = \sigma(w_x^{(\ell)} \boldsymbol{A}^T \boldsymbol{x} + b_x^{(\ell)}) \quad \text{if } \ell = 1 \qquad (9)$$

where $w_x^{(\ell)}, b_x^{(\ell)} \in \mathbb{R}$ are the per-layer weight and bias terms, respectively. $\sigma(\cdot)$ represents the nonlinear activation in each of the GNN layers, chosen here to be the hyperbolic tangent function.

2.3 Comparing Full and Reduced Models to Infer Granger Causality

Let $\Theta_y^{(full)}, \Theta_y^{(reduced)}, \Theta_x^{(full)} \in \mathbb{R}^{2L}$ denote the set of parameters for the full and reduced models, with $\Theta_x^{(full)} = \{w_x^{(1)}, ..., w_x^{(L)}; b_x^{(1)}, ..., b_x^{(L)}\}$ and $\Theta_y^{(full)}, \Theta_y^{(reduced)}$ defined analogously. We propose to jointly learn all the Θ parameters by minimizing the combined loss of the full and reduced models (Montalto et al., 2015):

$$\mathcal{L}^{(total)} = \sum_{x,y \in \mathcal{P}} \left(\sum_{v \in \mathcal{V}} \mathcal{L}(\hat{y}_v^{(full)}, y_v) + \sum_{v \in \mathcal{V}} \mathcal{L}(\hat{y}_v^{(reduced)}, y_v) \right) \tag{10}$$

where $\hat{y}_v^{(full)}$ and $\hat{y}_v^{(reduced)}$ correspond to the predictions for observation v by the full and reduced models, respectively. In a multivariate system, not all pairwise combinations of variables may be relevant: $\mathcal{P}$ is the subset of x, y variable pairs whose putative causal interactions are of interest.

To infer Granger causal interactions between y and x, we compare the set of loss terms associated with the full model $\mathcal{L}(\hat{y}_v^{(full)}, y_v)$ to those of the reduced model $\mathcal{L}(\hat{y}_v^{(reduced)}, y_v)$, noting that the precise functional form of $\mathcal{L}$ depends on the domain. If x does not Granger cause y, the full and reduced models will have similar explanatory power and hence similar loss values. Otherwise, the full model will have a lower loss. We assess this by a one-tailed Welch's t-test, where the alternative hypothesis is that the mean of the loss terms for the full model is less than the mean of the loss terms for the reduced model. We then score the interaction between each pair of variables as the $-\log_{10}(p)$-value associated with the Welch's t-test and rank the set of candidate x–y interactions by these scores, with a higher score indicating a more significant Granger causal interaction.

2.4 DOMAIN-SPECIFIC MODEL CUSTOMIZATION

The loss $\mathcal{L}$ should be chosen as appropriate for the domain. In our single-cell context, y corresponds to gene expression. We preprocessed RNA-seq transcript counts, normalizing and log transforming them so that the mean squared error loss became appropriate; we note that most methods based on Pearson correlation of gene expression also seek to minimize squared loss, explicitly or implicitly.

The model described above can also be customized to account for different degrees of lag. The number of GNN layers L corresponds to the maximum amount of past information desired. Similarly, to introduce a k-hop lag, the first k GNN layers would use the matrix $\boldsymbol{A}$ while the later layers would use $\boldsymbol{A}_+$. Note that in the above, we described a one-hop lag, which we use for all analyses.

2.5 TRAINING DETAILS AND HYPERPARAMETERS

GrID-Net models were trained using the Adam optimizer with a learning rate of 0.001 for 20 epochs or until convergence (defined to be the point at which the relative change in the loss function is less than $0.1/|\mathcal{P}|$ across consecutive epochs). A minibatch size of 1024 candidate peak–gene pairs was used during training, and trainable parameters in the model were initialized using Glorot initialization (Bengio & Glorot, 2010). All GrID-Net models consisted of $L = 10$ GNN layers. All models were implemented in PyTorch and trained on a single NVIDIA Tesla V100 GPU.

3 RELATED WORK

Single-cell multimodal methods: The emergence of single-cell multimodal assays has led to the development of tools to leverage the joint profiling of multiple genomic, epigenetic, and functional features in each cell. Many of these tools aim to synthesize the multiple modalities within a dataset to perform analysis tasks, such as dimensionality reduction, clustering, and visualization, that elucidate cell state or gene programs (Singh et al., 2021; Argelaguet et al., 2020; Gayoso et al., 2021). Meanwhile, efforts to detect associations at the level of specific features across modalities have thus far been limited. In particular, one of the primary opportunities made available by assays that profile chromatin accessibility and gene expression in the same cell is the identification of putative regulatory associations between accessible regions and genes. However, current approaches for inferring these associations predominantly rely on simply calculating correlations between these two modalities (Ma et al., 2020; Zhu et al., 2019; Liu et al., 2019). Not only do these correlations-based approaches fail to account for the temporal asynchrony between chromatin accessibility and gene expression, but they are also sensitive to the inherent noisiness of single-cell data (Zhang et al., 2019). These limitations point to the need for methods that go beyond analyzing mere correlations to leverage the joint dynamics of these features in causal inference frameworks, like Granger causality.

Granger causality: Recent approaches for extending Granger causality have primarily focused on enabling the detection of nonlinear interactions in multivariate time series data. One category

of methods involves nonlinear kernel-based regression models that infer linear Granger causality in a transformed feature space (Marinazzo et al., 2008; Ren et al., 2020). Neural network-based Granger causality models have also more recently been proposed to account for more generalized nonlinear interactions (Tank et al., 2018; Marcinkevičs & Vogt, 2021). Of these methods, efforts have been made to leverage specific neural network architectures to more fully take advantage of the sequential ordering of observations in time series data. Khanna & Tan (2020) proposed the economy statistical recurrent unit (eSRU) architecture to exploit the expressiveness of SRU units in a sample-efficient manner. Similarly, Temporal Causal Discovery Framework (TCDF) employs 1D-convolutional neural network layers as a means of capturing long-range temporal information along a single temporal axis (Nauta et al., 2019).

In our single-cell peak–gene inference application, performing Granger causal inference on a DAG (where each cell is a node and edges connect similar cells) allows us to capture the diversity of cell states as well as their dynamics. We note that using the standard Granger causality framework would require forcing this richly heterogeneous dataset into a sequential ordering of cells along pseudotime, a suboptimal choice due to the possibility of multiple differentiation trajectories, substantial biological and technical variation in single-cell data (Zhang et al., 2019), and multiple cells representing very similar cellular states (Baran et al., 2019).

4 RESULTS

Datasets and Preprocessing We analyzed three single-cell multimodal datasets with per-cell ATAC-seq and RNA-seq profiles that characterize a range of dynamical systems, including cell differentiation and drug-treatment responses (Cao et al., 2018; Chen et al., 2019; Ma et al., 2020). The **sci-CAR** dataset profiled 3,260 human lung adenocarcinoma-derived A549 cells after dexamethasone treatment. The **SNARE-seq** dataset evaluated 5,081 cells from the mouse neonatal cerebral cortex. Lastly, the **SHARE-seq** dataset contained 34,774 cells from differentiating skin cells in mouse. We applied geometric sketching (Hie et al., 2019) to identify a representative subset of 5,000 cells for the SHARE-seq dataset, which we use for all downstream analyses. These datasets are high-dimensional (about $200,000$ peaks and $30,000$ genes per cell) and sparse (over 99.5% of peaks in a cell have zero counts).

Following recommended practice (Hie et al., 2020), we preprocessed each dataset by applying $\log(1+\text{CPM}/10)$ and $\log(1+\text{CPM}/100)$ transformations to the raw RNA-seq and ATAC-seq count data, respectively; here CPM indicates counts per million. The transformed RNA-seq data for each gene was also divided by the maximum value for that gene. For each single-cell multimodal dataset, we then compiled a set of candidate peak–gene pairs (i.e., $\mathcal{P}$ in Eqn. 10) by selecting ATAC-seq peaks within 1 Mb from each gene (Appendix A.2). We chose this genomic-distance cutoff because almost all enhancer-gene pairs reported in a recent benchmark study were within 1 Mb (Moore et al., 2019). After this filtering, we evaluated 507,408, 916,939, and 872,039 candidate peak–gene pairs for the sci-CAR, SHARE-seq, and SNARE-seq datasets, respectively. Training GrID-Net on these sets of candidate peak–gene pairs took roughly 2–6 hours and required a maximum of 40 GB of RAM. Additional runtime and memory usage details can be found in Appendix A.7.

DAG construction For each of the three multimodal datasets, a kNN (k=15) graph was constructed on representations of cells that unified information from both gene expression and chromatin accessibility via Schema (Appendix A.1, Singh et al. (2021)). A pseudotime value was then inferred for each cell using diffusion pseudotime (Haghverdi et al., 2016), after which edges in the kNN graph are retained if they align with the direction of increasing pseudotime. Since each edge is oriented in the direction of increasing pseudotime, no cycles exist and the resulting graph is a DAG.

Evaluating causality by using independent information A key contribution of this work is a framework for systematically evaluating putative causal peak–gene relationships; to our knowledge, no such framework currently exists. Since ground truth genome-scale perturbational data is unavailable, we propose that the prediction of peak–gene associations be compared against independent, cell type-matched information sources. One information source we use are eQTLs from cell types matched to those profiled in the single-cell studies. An eQTL is a locus-gene pair that associates naturally occurring genetic variability at a locus to gene expression variations observed in large populations, thus serving as a proxy for direct perturbation of the locus (Nica & Dermitzakis, 2013). We

also evaluate against cell type-matched chromatin interaction (Hi-C/5C) data that report pairs of genomic loci that are spatially close in the 3D structure of the genome (Schoenfelder & Fraser, 2019). Intuitively, chromatin interaction and chromatin accessibility are both manifestations of the underlying gene regulatory process (**Figure** 1d): a non-coding region becomes accessible (ATAC-seq) for binding by a gene regulator and the chromatin 3D structure is arranged (chromatin interaction) so that the regulator is proximal to the target gene's promoter, thus enabling transcriptional control (Delaneau et al., 2019). As such, compared to random peak–gene pairs, causal pairs should more likely be in close spatial proximity.

We note that, in absolute terms, a high overlap of causal peak–gene pairs with these datasets should not be expected. Like ATAC-seq, both eQTLs and chromatin interaction capture the underlying biology only indirectly and are themselves prone to errors: eQTLs are confounded by correlated mutations elsewhere in the genome and Hi-C/5C data suffers from similar sparsity issues as ATAC-seq. Therefore, we focus here on a relative evaluation, comparing GrID-Net's overlap with these datasets against that of alternative approaches. Also, while good overlap with both eQTL and chromatin interaction datasets is desirable, we believe emphasis should be given to eQTLs since they are the better proxy for an actual perturbation.

Benchmarking We evaluated two correlation-based techniques used in previous single-cell multimodal studies (Cao et al., 2018; Chen et al., 2019; Ma et al., 2020): a) Pearson correlation (**Pearson Corr**) of peak counts and gene expression, with each cell considered an independent observation; and b) Pearson correlation computed after averaging the RNA-seq and ATAC-seq profiles of each cell and its 50 neighbors in the kNN graph of cells (**Pseudocell Corr**). The pseudocell approach, currently the state of the art, seeks robustness by replacing point estimates in the ATAC-seq (or RNA-seq) space with local averages. We use the absolute value of the correlation scores to rank peak–gene pairs in our analyses below. To assess the advantage of working with a partial ordering and operating on the kNN DAG, we also computed a traditional Granger causal estimate using vector autoregression (**VAR Granger**). There, data was forced into a total ordering along pseudotime by partitioning the pseudotime range into 100 bins of equal size and averaging the ATAC-seq and RNA-seq profiles across cells assigned to each bin (Appendix A.3).

Predicting eQTLs We obtained eQTL data from the GTEx project for the human cell types that most closely correspond to the cells from the single-cell multimodal assays (GTEx Consortium, 2013). For the lung adenocarcinoma-derived A549 cells in the sci-CAR dataset, we used eQTL data for human lungs. Meanwhile, for the mouse skin and neonatal cortex cells from the SHARE-seq and SNARE-seq datasets, respectively, we used eQTL data for human skin and the human cerebral cortex, mapping them to the mouse genome using *liftOver* (Gonzalez et al., 2021). For each dataset, we retained peak–gene pairs with a matching locus–gene pair in the eQTL data (Appendix A.4).

In order to assess the effectiveness of GrID-Net in predicting eQTLs, we labeled peak–gene pairs associated with eQTLs as *true* (eQTL $p < 10^{-10}$) or *false* (eQTL $p > 0.9$), discarding pairs that were not in either category. GrID-Net and the alternative methods were evaluated on their accuracy in predicting the *true* eQTLs based on the ranking of association scores generated by each method for each peak–gene pair. Across all single-cell datasets, GrID-Net dramatically outperformed the other methods in predicting eQTLs, achieving the highest AUPRC (area under precision-recall curve) and AUROC (area under receiver operating characteristic) in all cases, indicating that peak–gene pairs prioritized by GrID-Net more closely align with evidence from population genetics (Table 1). We also found the two correlation approaches to have similar predictive power to each other, suggesting that simply averaging values in a local neighborhood is not sufficient. The VAR Granger method performed substantially worse than GrID-Net as well, which points to the advantage of DAG-based causal inference for such systems over standard causal inference on totally ordered data.

Table 1: eQTL prediction accuracy

Method	sci-CAR		SHARE-seq		SNARE-seq	
	AUPRC	AUROC	AUPRC	AUROC	AUPRC	AUROC
Pearson Corr	0.0594	0.514	0.0317	0.505	0.0096	0.619
Pseudocell Corr	0.0591	0.512	0.0165	0.554	0.0084	0.573
VAR Granger	0.0542	0.493	0.0321	0.480	0.0037	0.439
GrID-Net	**0.0697**	**0.541**	**0.0529**	**0.590**	**0.0164**	**0.634**

Predicting chromatin interactions We obtained chromatin interaction data for the specific cell types and experimental settings profiled in the three single-cell multimodal datasets (D'Ippolito et al., 2018; Dixon et al., 2012; Poterlowicz et al., 2017). Chromatin interactions were determined for a peak–gene pair by identifying the pairs of genomic windows in the corresponding chromatin interaction dataset that overlapped the peak and the transcription start site (TSS) of the gene. We then identified the peak–gene pairs with significant chromatin interactions (Appendix A.5) and aimed to compare the accuracy of GrID-Net and alternative methods in predicting these pairs. We found that GrID-Net achieved the highest AUPRC values in this prediction task across all datasets, demonstrating its ability to better identify these key functional features in gene regulation (Table 2). It also achieved the highest AUROC values in all but the sci-CAR dataset. As a measure, AUROC is more susceptible to class imbalances than AUPRC and this may explain GrID-Net's muted performance on that metric in the sci-CAR data.

Table 2: Chromatin interaction prediction accuracy

Method	sci-CAR		SHARE-seq		SNARE-seq	
	AUPRC	AUROC	AUPRC	AUROC	AUPRC	AUROC
Pearson Corr	0.0093	0.519	0.0126	0.481	0.0106	0.513
Pseudocell Corr	0.0095	**0.522**	0.0117	0.475	0.0108	0.517
VAR Granger	0.0084	0.506	0.0207	0.474	0.0102	0.504
GrID-Net	**0.0098**	0.508	**0.0653**	**0.862**	**0.0113**	**0.542**

GrID-Net is more robust to sparse data than existing approaches With the sparsity of single-cell multimodal data expected to remain a concern, we assessed GrID-Net's robustness to such sparsity. Focusing on the subset of genes or peaks where data is especially sparse (fewer than 5% or 2% of cells for genes or peaks, respectively), we evaluated it on the eQTL prediction task, finding that its outperformance compared to the alternatives was accentuated in this subset (Table 3). GrID-Net also noticeably outperforms the other methods in predicting significant chromatin interactions associated with peak–gene pairs consisting of sparse genes or peaks. These results provide further evidence of GrID-Net's utility in extracting insights from sparse single-cell data.

Table 3: eQTL and chromatin interaction prediction accuracy for sparse genes and peak (AUPRC)

Method	eQTL						Chromatin interactions					
	sci-CAR		SHARE-seq		SNARE-seq		sci-CAR		SHARE-seq		SNARE-seq	
	Sparse Genes	Sparse Peaks	Sparse Genes	Sparse Peaks	Sparse Genes	Sparse Peaks	Sparse Genes	Sparse Peaks	Sparse Genes	Sparse Peaks	Sparse Genes	Sparse Peaks
Pearson Corr	0.0607	0.0566	0.0438	0.0330	0.0136	0.0078	0.0083	0.0084	0.0111	0.0133	0.0104	0.0088
Pseudocell Corr	0.0589	0.0562	0.0540	0.0224	0.0122	0.0058	0.0083	0.0086	0.0116	0.0099	0.0103	0.0089
VAR Granger	0.0494	0.0520	0.0444	0.0168	0.0045	0.0036	0.0075	0.0081	0.0049	0.0664	0.0095	0.0087
GrID-Net	**0.0706**	**0.0676**	**0.0764**	**0.1092**	**0.0193**	**0.0162**	**0.0096**	**0.0087**	**0.0646**	**0.0678**	**0.0114**	**0.0095**

Peak–gene pairs detected by GrID-Net are enriched for TF binding motifs To assess the functional relevance of the detected peak–gene pairs, we aimed to relate these peaks to the transcription factors (TFs) that might bind there. Accordingly, we selected the highest scoring 1% of peak–gene pairs from each of GrID-Net and its alternatives. We then used Homer (Heinz et al., 2010) to identify TF binding motifs that are co-enriched in sets of peaks linked to a particular gene (Appendix A.6). The presence of these co-enriched motifs is potentially indicative of coordinated regulatory control, whereby a gene is modulated by multiple TF binding events featuring the same or related sets of TFs (Pott & Lieb, 2015; Carleton et al., 2017). Consequently, the co-occurrence of these motifs serves as a proxy for the functional significance of the proposed peak–gene pairs. Across all three datasets, we observed that the peak–gene pairs detected by GrID-Net were associated with 10–50 times more enriched TF binding motifs than the other methods (Table 4). These results suggest that GrID-Net may serve a valuable role in unveiling mechanistic insights into TF-gene regulatory control.

Table 4: Number of putative TF-target gene relations detected from peak–gene pairs

	sci-CAR	SHARE-seq	SNARE-seq
Pearson Corr	2	95	0
Pseudocell Corr	16	68	23
VAR Granger	2	45	56
GrID-Net	**92**	**1375**	**3037**

GrID-Net reveals link between genomic distance and temporal regulatory control GrID-Net's explicit consideration of the temporal asynchrony between chromatin accessibility and gene expression enables the genome-wide study of lags in gene regulation for the very first time. We classified peak–gene pairs as *proximal* (genomic distance < 10kb) or *distal* (≥ 10 kb). We then evaluated different architectures of GrID-Net that vary in the number of GNN layers L, hypothesizing that architectures with higher L (i.e., the degree of lag between peaks and genes that can be captured) may detect more distal peak–gene pairs. Inspecting the top 1% of peak–gene pairs detected by different architectures, we found this to indeed be the case: across all datasets, architectures with higher L detected smaller proportions of proximal peak–gene pairs (Table 5), suggesting that distal peak–gene interactions may be marked by longer lags between peaks and genes. This finding indicates, to our knowledge, a previously unreported link between genomic distance and the lag between a regulatory element's accessibility and its target gene's expression. One hypothesis is that the complex coordination of regulatory factors required to enable distal regulatory interactions is associated with greater lags between distal regulatory elements' accessibility and the gene expression changes they affect (Panigrahi & O'Malley, 2021). GrID-Net, thus, presents a novel result that provides insight into the temporal dynamics governing distal regulatory elements and their gene expression control.

Table 5: Percentages of peak–gene pairs within 10kb

	sci-CAR	SHARE-seq	SNARE-seq
2 GNN Layers	14.5%	6.04%	2.53%
5 GNN Layers	14.1%	5.20%	2.35%
10 GNN Layers	13.7%	5.33%	2.25%

5 DISCUSSION

We extend the applicability of Granger causal analysis to dynamical systems represented as DAGs, introducing GrID-Net to perform such inference. To best leverage the dynamics expressed in these DAG-based systems, GrID-Net takes advantage of the expressive power of graph neural networks, enabling the detection of long-range nonlinear interactions between variables. We focused on using GrID-Net to study multimodal single-cell gene regulatory dynamics, given the prevalence of graph-based representations of cellular landscapes in the field. GrID-Net demonstrated substantial improvements in accuracy relative to current methods in predicting peak–gene links from independent, cell type-matched information sources. GrID-Net exhibited particularly strong performance in predicting links involving specific peaks and genes for which the data was sparse, a feat indicative of its value for extracting meaningful insights from sparse single-cell data.

GrID-Net's overall improvement in accuracy for the eQTL and chromatin interaction prediction tasks is also noteworthy, as it reveals important advantages that can be unlocked by using GrID-Net for studying gene regulation in future work. The large population sizes needed for eQTL studies and the relatively low resolution of chromatin interaction data (e.g., tens of kilobases in Hi-C) currently act as major limitations for exploring the space of genomic loci-gene interactions. With the higher resolution of ATAC peaks (~ 1 kb), regulatory links detected by GrID-Net from single-cell multimodal assays may be used to more precisely hone in on functional regions within noncoding loci or even be used to complement existing chromatin interaction data to aid in this goal (Fulco et al., 2019; Nasser et al., 2021). This increased resolution, combined with the genome-wide scale of single-cell multimodal assays, positions GrID-Net to serve as a vital tool for furthering our understanding of key aspects of gene regulatory dynamics relevant to fundamental biology and disease.

Lastly, the ever growing quantities of network data across numerous non-biological domains, including social media networks and financial transaction networks, also points to the broader applicability of GrID-Net. The ability of GrID-Net to capture long-range, nonlinear interactions in DAGs opens the door to novel analyses on the many existing datasets that are characterized by these graph-based dynamics as well.

REPRODUCIBILITY STATEMENT

In order to ensure that the full set of results presented in this paper are reproducible, we have included source code for our method, which can be found in the supplementary materials. In addition,

all analyses described in the paper were performed on publicly available datasets, for which we have provided references throughout the main text. Similarly, the software tools used for biology-specific analyses (e.g., Homer) are publicly available, and we have described the specific settings that were used for this study in the "Appendix" section. We have also detailed pre-processing steps for the various single-cell datasets in the main text and included both pre-processing and analysis procedures related to the eQTL, Hi-C/5C, and TF binding motif data in the "Appendix" section. We also made sure to provide details on the specific training details and hyperparameter choices related to training our neural network model in a dedicated subsection in the "Methods" section entitled "Training Details and Hyperarameters".

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A   APPENDIX

A.1   DAG CONSTRUCTION

For the three single-cell multimodal datasets, we applied a standard $\log(1+\text{CPM}/10)$ transformation to the raw RNA-seq counts data and a $\log(1+\text{CPM}/100)$ transformation to the raw ATAC-seq counts data, where CPM indicates counts per million. Genes and accessible regions detected in fewer than 0.1% of cells in each dataset were excluded. For each dataset, the expression of each gene was scaled to unit variance and zero mean, after which the top 2000 most highly variable genes (Satija et al., 2015) were retained. In the original SHARE-seq and SNARE-seq studies, topic modeling-based approaches were used to transform ATAC-seq counts data for downstream trajectory analyses, prompting us to transform the ATAC-seq counts in these datasets using *tf-idf*, a procedure derived from text-based topic modeling that has been applied extensively to analyze single-cell ATAC-seq data (Cusanovich et al., b;c;a).

The normalized RNA-seq and ATAC-seq data were then projected onto their top 100 principal com-ponents (excluding components that have Spearman correlation $\rho > 0.9$ with the total counts per cell). Schema (Singh et al., 2021) was then used to generate unified representations for each cell that synthesize information from both the RNA-seq and ATAC-seq representations. For the sci-CAR and SNARE-seq datasets, RNA-seq was used as the primary modality in Schema with ATAC-seq acting as a secondary modality. For the sci-CAR dataset, the time of data collection post-drug treat-ment was used as a tertiary modality. Meanwhile, ATAC-seq served as the primary modality for the SHARE-seq dataset, as the original SHARE-seq study identified strong cell cycle effects rep-resented in the RNA-seq data for certain subpopulations of cells and consequently used ATAC-seq as the modality for analyzing lineage dynamics for the dataset (Ma et al., 2020). We set the mini-mum correlation between the primary and the remaining modalities to be 0.9 in Schema. For each dataset, a kNN graph ($k = 15$) based on distances between the unified representations generated by Schema was constructed. Pseudotime stamps were then inferred for each cell based on this kNN graph (Haghverdi et al., 2016). Finally, edges in the kNN graph that cohere with the direction of increasing pseudotime are retained to generate a DAG.

A.2 SINGLE-CELL DATA PREPROCESSING AND CANDIDATE PEAK–GENE PAIR SELECTION

We applied the same CPM and log transformations as above to the raw RNA-seq and ATAC-seq counts data for all three single-cell multimodal datasets. We then scaled the RNA-seq data for each gene by the maximum value for that gene, so that expression values for all genes in each dataset are in the same [0,1] range. The peak–gene pairs considered for each dataset were determined by selecting genes and peaks detected in greater than 0.1% of cells and then retaining all peak–gene pair combinations in which the peak is within 1 Mb of the gene. For the SHARE-seq dataset, we only considered peaks detected in greater than 1% of cells so as to have a comparable set of candidate peak–gene pairs across the three datsets.

A.3 BENCHMARKING

For the VAR Granger model, peak–gene pairs were scored on the $-\log_{10}(p)$-value associated with the statistical significance of Granger causality, as calculated using the `grangercausalitytests` function with default settings in the `statsmodels` package (Seabold & Perktold, 2010).

A.4 MATCHING EQTLS WITH SINGLE-CELL ATAC-SEQ & RNA-SEQ DATA

The procedure to match a peak–gene pair with a eQTL variant–gene pair involved identifying all catalogued variants positioned in the peak of interest and selecting the variant with the most statistically significant association with the gene in the peak–gene pair. To perform this procedure for the SNARE-seq and SHARE-seq datasets, we used liftOver (Gonzalez et al., 2021) with default settings to convert human genomic coordinates represented in the GTEx datasets to mouse coordinates.

A.5 CHROMATIN INTERACTION DATA PREPROCESSING

For the SHARE-seq and SNARE-seq datasets, the corresponding chromatin interaction datasets contained scores indicating the frequency of interactions between genomic windows. We retained peak–gene pairs that were matched with the genomic windows presented in the chromatin interaction datasets and chose the most frequent 1% of interactions as the set to be predicted. The chromatin interaction data associated with the sci-CAR dataset already included a set of statistically significant chromatin interactions across the entire genome, so the prediction task for the sci-car dataset involved predicting this set of interactions.

A.6 TF BINDING MOTIF ENRICHMENT ANALYSIS

For each gene that was associated with more than 5 peaks in the highest scoring 1% peak–gene pairs, we applied HOMER with default parameters on the genomic windows defined by the set of peaks to test for enrichment of TF binding motifs (Heinz et al., 2010).

A.7 RUNTIME AND MEMORY USAGE

GrID-Net's runtime and memory usage are dependent on a variety of factors, most notably the number of nodes in the DAG describing the system of interest. In our model, the DAG is represented by the $N \times N$ adjacency matrices $\boldsymbol{A}$ and $\boldsymbol{A}_+$, where n is the number of nodes in the DAG (i.e. number of cells in a multimodal single-cell dataset). These matrices serve as the primary bottlenecks for both runtime and memory usage. Memory usage scales quadratically with n, while calculations involving $\boldsymbol{A}$ or $\boldsymbol{A}_+$ can be parallelized on a GPU to reduce the runtime. In addition, training time scales linearly with the number of candidate peak-gene pairs. In our results above, we evaluated 507,408, 916,939, and 872,039 candidate peak–gene pairs for the sci-CAR, SHARE-seq, and SNARE-seq datasets, respectively. For these same datasets, we considered systems of 3,260, 5,000, and 5,081 cells. The average total runtime for GrID-Net was 1.8 hours, 5.5 hours, and 5.3 hours, while peak memory usage was approximately 31 GB, 35 GB, and 39 GB for these respective datasets.