Robustifying Algorithms of Learning Latent Trees with Vector Variables

Abstract

We consider learning the structures of Gaussian latent tree models with vector observations when a subset of them are arbitrarily corrupted. First, we present the sample complexities of Recursive Grouping (RG) and Chow-Liu Recursive Grouping (CLRG) without the assumption that the effective depth is bounded in the number of observed nodes, significantly generalizing the results in Choi et al. (2011). We show that Chow-Liu initialization in CLRG greatly reduces the sample complexity of RG from being exponential in the diameter of the tree to only logarithmic in the diameter for the hidden Markov model (HMM). Second, we robustify RG, CLRG, Neighbor Joining (NJ) and Spectral NJ (SNJ) by using the truncated inner product. These robustified algorithms can tolerate a number of corruptions up to the square root of the number of clean samples. Finally, we derive the first known instance-dependent impossibility result for structure learning of latent trees. The optimalities of the robust version of CLRG and NJ are verified by comparing their sample complexities and the impossibility result.

1 Introduction

Latent graphical models provide a succinct representation of the dependencies among observed and latent variables. Each node in the graphical model represents a random variable or a random vector, and the dependencies among these variables are captured by the edges among nodes. Graphical models are widely used in domains from biology [1], computer vision [2] and social networks [3].

This paper focuses on the structure learning of latent tree Gaussian graphical models (GGM) in which the node observations are random *vectors* and a subset of the observations can be *arbitrarily corrupted*. This classical problem, in which the variables are *clean scalar* random variables, has been studied extensively in the past decades. The first information distance-based method, NJ, was proposed in [1] to learn the structure of phylogenetic trees. This method makes use of additive information distances to deduce the existence of hidden nodes and introduce edges between hidden and observed nodes. RG, proposed in [4], generalizes the information distance-based methods to make it applicable for the latent graphical models with general structures. Different from these information distance-based methods, quartet-based methods [5] utilize the relative geometry of every four nodes to estimate the structure of the whole graph. Although experimental comparisons of these algorithms were conducted in some works [4, 6, 7], since there is no instance-dependent impossibility result of the sample complexity of structure learning of latent tree graphical models, no thorough theoretical comparisons have been made, and the optimal dependencies on the diameter of graphs and the maximal distance between nodes $\rho_{\max}$ have not been found.

The success of the previously-mentioned algorithms relies on the assumption that the observations are i.i.d. samples from the generating distribution. The structure learning of latent graphical models in presence of (random or adversarial) noise remains a relatively unexplored problem. There are some

works studying the problem of structure learning of graphical models with noisy samples, where all the nodes in the graphical models are observed and not hidden. Several assumptions on the additive noise are made in these works, which limit the use of these proposed algorithms. For example, the covariance matrix of the noise is specified in [8], the independence and/or distribution of the noise is assumed in [9–11, 7]. In contrast, we consider the structure learning of latent tree graphical models with *arbitrary* corruptions, where the boundness and independence of noise are not required [12]. Furthermore, the corruptions are allowed to be presented at any position in the data; they do not appear solely as outliers. In this work, we derive a bound on the maximum number of corruptions that can be tolerated and yet structure learning can succeed with high probability.

Firstly, we derive the sample complexities of RG and CLRG where each node represents a random *vector*; this differs from previous works where each node is *scalar* random variable (e.g., [4, 13]). We explore the dependence of the sample complexities on the parameters. The vanilla CLRG takes in the mutual informations and builds a max-weight spanning tree to learn the tree. However, for latent trees, we work with information distances and the mutual information is, in general, not monotone in the distance (unless the variables are scalar). To implement CLRG on GGMs with vector variables, we consider a class of GGMs in which parents and children nodes are connected by linear Gaussian channels. Sufficient conditions are derived to ensure that the mutual information is monotone in the distance, ensuring that CLRG is correctly implemented. Our sample complexity analysis proves the effectiveness of the Chow-Liu initialization in CLRG; this has been only verified experimentally [4]. In the HMM, we show that the Chow-Liu initialization reduces the sample complexity of RG which is $O\big((\frac{9}{2})^{\mathrm{Diam}(\mathbb{T})}\big)$ to $O\big(\log \mathrm{Diam}(\mathbb{T})\big)$, where $\mathrm{Diam}(\mathbb{T})$ is the tree diameter.

Secondly, we robustify RG, CLRG, NJ and SNJ by using the truncated inner product [14] to estimate the information distances in the presence of arbitrary corruptions. We derive their sample complexities and show that they can tolerate $n_1 = O\big(\frac{\sqrt{n_2}}{\log n_2}\big)$ corruptions, where n_2 is the number of clean samples.

Finally, we derive the first known instance-dependent impossibility result for learning latent trees. The dependencies on the number of observed nodes and the maximum distance $\rho_{\max}$ are delineated. The comparison of the sample complexities of the structure learning algorithms and the impossibility result demonstrates the optimality of Robust Chow-Liu Recursive Grouping (RCLRG) and Robust Neighbor Joining (RNJ) in $\mathrm{Diam}(\mathbb{T})$ for some archetypal latent tree structures.

Notation We use san-serif letters x, boldface letters $\mathbf{x}$, and bold uppercase letters $\mathbf{X}$ to denote variables, vectors and matrices, respectively. The notations $[\mathbf{x}]_i$, $[\mathbf{X}]_{ij}$, $[\mathbf{X}]_{:,j}$ and $\mathrm{diag}(\mathbf{X})$ are respectively the ith entry of vector $\mathbf{x}$, the $(i,j)^{\mathrm{th}}$ of entry $\mathbf{X}$, the jth column of $\mathbf{X}$, and the diagonal entries of matrix $\mathbf{X}$. The notation $x^{(k)}$ represents the kth sample of x. For a tree $\mathbb{T} = (\mathcal{V}, \mathcal{E})$, the internal (non-leaf) nodes, the maximal degree and the diameter of $\mathbb{T}$ are denoted as $\mathrm{Int}(\mathbb{T})$, $\mathrm{Deg}(\mathbb{T})$ and $\mathrm{Diam}(\mathbb{T})$, respectively. We denote the closed neighborhood and the degree of x_i as $\mathrm{nbd}[x_i; \mathbb{T}]$ and $\deg(i)$, respectively. The length of the path connecting x_i and x_j is denoted as $\mathrm{d}_{\mathbb{T}}(x_i, x_j)$.

2 Preliminaries and problem statement

A GGM [15] is a multivariate Gaussian distribution that factorizes according to an undirected graph $\mathbb{G} = (\mathcal{V}, \mathcal{E})$. More precisely, a l_{sum}-dimensional random vector $\mathbf{x} = [\mathbf{x}_1, \ldots, \mathbf{x}_d]^{\mathrm{T}}$, where $\mathbf{x}_i \in \mathbb{R}^{l_i}$ and $l_{\mathrm{sum}} = \sum_{i=1}^d l_i$, follows a Gaussian distribution $\mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$, and it is said to be *Markov* on a graph $\mathbb{G} = (\mathcal{V}, \mathcal{E})$ with vertex set $\mathcal{V} = \{x_1, \ldots, x_d\}$ and edge set $\mathcal{E} \subseteq \binom{\mathcal{V}}{2}$ and $(x_i, x_j) \in \mathcal{E}$ if and only if the $(i,j)^{\mathrm{th}}$ block $\boldsymbol{\Theta}_{ij}$ of the precision $\boldsymbol{\Theta} = \boldsymbol{\Sigma}^{-1}$ is not the zero matrix $\mathbf{0}$. We focus on tree-structured graphical models, which factorize according to acyclic and connected (tree) graphs.

A special class of graphical models is the set of *latent* graphical models $\mathbb{G} = (\mathcal{V}, \mathcal{E})$. The vertex set $\mathcal{V}$ is decomposed as $\mathcal{V} = \mathcal{V}_{\mathrm{hid}} \cup \mathcal{V}_{\mathrm{obs}}$. We only have access to n i.i.d. samples drawn from the observed set of nodes $\mathcal{V}_{\mathrm{obs}}$. The two goals of any structure learning algorithm are to learn the identities of the hidden nodes $\mathcal{V}_{\mathrm{hid}}$ and how they are connected to the observed nodes.

2.1 System model for arbitrary corruptions

We consider tree-structured GGMs $\mathbb{T} = (\mathcal{V}, \mathcal{E})$ with observed nodes $\mathcal{V}_{\mathrm{obs}} = \{x_1, \cdots, x_o\}$ and hidden nodes $\mathcal{V}_{\mathrm{hid}} = \{x_{o+1}, \cdots, x_{o+h}\}$, where $\mathcal{V} = \mathcal{V}_{\mathrm{hid}} \cup \mathcal{V}_{\mathrm{obs}}$ and $\mathcal{E} \subseteq \binom{\mathcal{V}}{2}$. Each node x_i

represents a random *vector* $\mathbf{x}_i \in \mathbb{R}^{l_i}$. The concatenation of these random vectors is a multivariate Gaussian random vector with zero mean and covariance matrix $\mathbf{\Sigma}$ with size $l_{\text{sum}} \times l_{\text{sum}}$.

We have n i.i.d. samples $\mathbf{X}_j = [\mathbf{x}_1^{(j)\text{T}}, \cdots, \mathbf{x}_o^{(j)\text{T}}]^\text{T} \in \mathbb{R}^{l_{\text{sum}}}, j = 1, \ldots, n$ drawn from the observed nodes $\mathcal{V}_{\text{obs}} = \{x_1, \cdots, x_o\}$. Furthermore, the data matrix $\mathbf{X}_1^n = [\mathbf{X}_1, \cdots, \mathbf{X}_n]^\text{T} \in \mathbb{R}^{n \times l_{\text{sum}}}$ may contain some corrupted elements. We allow $n_1/2$ samples of a variable to be arbitrarily corrupted, which means that there are at most $n_1/2$ corrupted terms in each column of $\mathbf{X}_1^n$, and the remaining $n - n_1/2$ samples in this column are clean. In particular, the corrupted samples in different columns need not be in the same rows. If the corruptions in different columns lie in the same rows, as shown in (the left of) Fig. 3, all the samples in the corresponding rows are corrupted; these are called *outliers*. Obviously, outliers form a special case of our corruption model. Since each variable has at most $n_1/2$ corrupted samples, the sample-wise inner product between two variables has at least $n_2 = n - n_1$ clean samples. There is no constraint on the statistical dependence or patterns of the corruptions. Unlike fixing the covariance matrix of the noise [8] or keeping the noise independent [9], we allow *arbitrary* corruptions on the samples, which means that the noise can have unbounded amplitude, can be dependent, and even can be generated from another graphical model (as we will see in the experimental results in Section 3.6).

2.2 Structural and distributional assumptions

To construct the correct latent tree from samples of observed nodes, it is imperative to constrain the class of latent trees to guarantee that the information from the distribution of observed nodes $p(\mathbf{x}_1, \ldots, \mathbf{x}_o)$ is sufficient to construct the tree. The distribution $p(\mathbf{x}_1, \ldots, \mathbf{x}_{o+h})$ of the observed and hidden nodes is said to have a *redundant* hidden node x_j if the distribution observed nodes $p(\mathbf{x}_1, \ldots, \mathbf{x}_o)$ remains the same after we marginalize over x_j. To ensure that a latent tree can be constructed with no ambiguity, we need to guarantee that the true distribution does not have any redundant hidden node(s), which is achieved by following two conditions [16]: (C1) Each hidden node has at least three neighbors; the set of such latent trees is denoted as $\mathcal{T}_{\geq 3}$; (C2) Any two variables connected by an edge are neither perfectly dependent nor independent.

Assumption 1. *The dimensions of all the random vectors are all equal to $l_{\max}$.*

In fact, we only require the random vectors of the internal (non-leaf) nodes to have the same length. However, for ease of notation, we assume that the dimensions of all random vectors are $l_{\max}$.

Assumption 2. *For every $x_i, x_j \in \mathcal{V}$, the covariance matrix $\mathbf{\Sigma}_{ij} = \mathbb{E}\big[\mathbf{x}_i \mathbf{x}_j^\text{T}\big]$ has full rank, and the smallest singular value of $\mathbf{\Sigma}_{ij}$ is lower bounded by $\gamma_{\min}$, i.e.,*

$$\sigma_{l_{\max}}(\mathbf{\Sigma}_{ij}) \geq \gamma_{\min} \quad \text{for all} \quad x_i, x_j \in \mathcal{V}, \tag{1}$$

where $\sigma_i(\mathbf{\Sigma})$ is the ith largest singular value of $\mathbf{\Sigma}$.

This assumption is a strengthening of Condition (C2) when each node represents a random vector.

Assumption 3. *The determinant of the covariance matrix of any node $\mathbf{\Sigma}_{ii} = \mathbb{E}\big[\mathbf{x}_i \mathbf{x}_i^\text{T}\big]$ is lower bounded by $\delta_{\min}$, and the diagonal terms of the covariance matrix are upper bounded by $\sigma_{\max}^2$, i.e.,*

$$\min_{x_i \in \mathcal{V}} \det(\mathbf{\Sigma}_{ii}) \geq \delta_{\min} \quad \text{and} \quad \max_{x_i \in \mathcal{V}} \text{diag}(\mathbf{\Sigma}_{ii}) \leq \sigma_{\max}^2. \tag{2}$$

Assumption 3 is natural; otherwise, $\mathbf{\Sigma}_{ii}$ may be arbitrarily close to a singular matrix.

Assumption 4. *The degree of each node is upper bounded by $d_{\max}$, i.e., $\text{Deg}(\mathbb{T}) \leq d_{\max}$.*

2.3 Information distance

We define the *information distance* for Gaussian random vectors and prove that it is additive for trees.

Definition 1. *The information distance between nodes x_i and x_j is*

$$\text{d}(x_i, x_j) = -\log \frac{\prod_{k=1}^{l_{\max}} \sigma_k(\mathbf{\Sigma}_{ij})}{\sqrt{\det(\mathbf{\Sigma}_{ii})\det(\mathbf{\Sigma}_{jj})}}. \tag{3}$$

Condition (C2) can be equivalently restated as constraints on the information distance.

Assumption 5. *There exist two constants $0 < \rho_{\min} \leq \rho_{\max} < \infty$ such that.*

$$\rho_{\min} \leq d(x_i, x_j) \leq \rho_{\max} \quad \textit{for all} \quad x_i, x_j \in \mathcal{V}. \tag{4}$$

Assumptions 2 and 5 both describe the properties of the correlation between random vectors from different perspectives. In fact, we can relate the constraints in these two assumptions as follows:

$$\gamma_{\min} e^{\rho_{\max}/l_{\max}} \geq \delta_{\min}^{1/l_{\max}}. \tag{5}$$

Proposition 1. *If Assumptions 1 and 2 hold, $d(\cdot, \cdot)$ defined in Definition 1 is additive on the tree-structured GGM $\mathbb{T} = (\mathcal{V}, \mathcal{E})$. In other words, $d(x_i, x_k) = d(x_i, x_j) + d(x_j, x_k)$ holds for any two nodes $x_i, x_k \in \mathcal{V}$ and any node x_j on the path connecting x_i and x_k in $\mathbb{T}$.*

This additivity property is used extensively in the following algorithms. It was first stated and proved in Huang et al. [17]. We provide an alternative proof in Appendix G.

3 Robustifying latent tree structure learning algorithms

3.1 Robust estimation of information distances

Before delving into the details of robustifying latent tree structure learning algorithms, we first introduce the truncated inner product [14], which estimates the correlation against arbitrary corruption effectively and serves as a basis for the robust latent tree structure learning algorithms. Given $\mathbf{a}, \mathbf{b} \in \mathbb{R}^n$ and an integer n_1, we compute $q_i = a_i b_i$ for $i = 1, 2, \ldots, n$ and sort $\{|q_i|\}$. Let Υ be the index set of the $n - n_1$ smallest $|q_i|$'s. The truncated inner product is $\langle \mathbf{a}, \mathbf{b} \rangle_{n_1} = \sum_{i \in \Upsilon} q_i$. Note that the implementation of the truncated inner product requires the knowledge of corruption level n_1.

To estimate the information distance defined in Definition 1, we implement the truncated inner product to estimate each term of $\mathbf{\Sigma}_{ij}$, i.e., $[\hat{\mathbf{\Sigma}}_{ij}]_{st} = \frac{1}{n-n_1} \langle [\mathbf{X}_1^n]_{:, (i-1)l_{\max}+s}, [\mathbf{X}_1^n]_{:, (j-1)l_{\max}+t} \rangle_{n_1}$. Then the information distance is computed based on this estimate of $\mathbf{\Sigma}_{ij}$ as

$$\hat{d}(x_i, x_j) = -\log \prod_{k=1}^{l_{\max}} \sigma_k(\hat{\mathbf{\Sigma}}_{ij}) + \frac{1}{2} \log \det(\hat{\mathbf{\Sigma}}_{ii}) + \frac{1}{2} \log \det(\hat{\mathbf{\Sigma}}_{jj}). \tag{6}$$

The truncated inner product guarantees that $\hat{\mathbf{\Sigma}}_{ij}$ converges in probability to $\mathbf{\Sigma}_{ij}$, which further ensures the convergence of the singular values and the determinant of $\mathbf{\Sigma}_{ij}$ to their nominal values.

Proposition 2. *If Assumptions 1 and 2 hold, the estimate of the information distance between x_i and x_j based on the truncated inner product $\hat{d}(x_i, x_j)$ satisfies*

$$\mathbb{P}\left(\left| \hat{d}(x_i, x_j) - d(x_i, x_j) \right| > \frac{2l_{\max}^2}{\gamma_{\min}}(t_1 + t_2) \right) \leq 2l_{\max}^2 e^{-\frac{3n_2}{16\kappa n_1} t_1} + l_{\max}^2 e^{-c \frac{n_2}{\kappa^2} t_2^2}, \tag{7}$$

where $t_2 < \kappa = \max\{\sigma_{\max}^2, \rho_{\min}\}$, and c is an absolute constant.

The first and second parts of (7) originate from the corrupted and clean samples respectively.

3.2 Robust Recursive Grouping algorithm

The RG algorithm was proposed by [4] to learn latent tree models with additive information distances. We extend the RG to be applicable to GGMs with vector observations and robustify it to learn the tree structure against arbitrary corruptions. We call this robustified algorithm Robust Recursive Grouping (RRG). RRG makes use of the additivity of information distance to identify the relationship between nodes. For any three nodes x_i, x_j and x_k, the difference between the information distances $d(x_i, x_k)$ and $d(x_j, x_k)$ is denoted as $\Phi_{ijk} = d(x_i, x_k) - d(x_j, x_k)$.

Lemma 3. *[4] For information distances $d(x_i, x_j)$ for all nodes $x_i, x_j \in \mathcal{V}$ in a tree $\mathbb{T} \in \mathcal{T}_{\geq 3}$, Φ_{ijk} has following two properties: (1) $\Phi_{ijk} = d(x_i, x_j)$ for all $x_k \in \mathcal{V}\backslash\{x_i, x_j\}$ if and only if x_j is a leaf node and x_i is the parent of x_j and (2) $-d(x_i, x_j) < \Phi_{ijk'} = \Phi_{ijk} < d(x_i, x_j)$ for all $x_k, x_{k'} \in \mathcal{V}\backslash\{x_i, x_j\}$ if and only if x_i and x_j are leaves and share the same parent.*

RRG initializes the active set Γ^1 to be the set of all observed nodes. In the ith iteration, as shown in Algorithm 1, RRG adopts Lemma 3 to identify relationships among nodes in active set Γ^i, and it removes nodes identified as siblings from Γ^i and adds newly introduced hidden nodes to form the active set Γ^{i+1} in the $(i+1)^{\text{st}}$ iteration. The procedure of estimating the distances between the newly-introduced hidden node x_{new} and other nodes is as follows. For the node x_i which is the child of x_{new}, i.e., $x_i \in \mathcal{C}(x_{\text{new}})$, the information distance is estimated as

$$\hat{\mathrm{d}}(x_i, x_{\text{new}}) = \frac{1}{2(|\mathcal{C}(x_{\text{new}})| - 1)} \left(\sum_{j \in \mathcal{C}(x_{\text{new}})} \hat{\mathrm{d}}(x_i, x_j) + \frac{1}{|\mathcal{K}_{ij}|} \sum_{k \in \mathcal{K}_{ij}} \hat{\Phi}_{ijk} \right), \tag{8}$$

where $\mathcal{K}_{ij} = \big\{ x_k \in \mathcal{V} \backslash \{x_i, x_j\} : \max \big\{ \hat{\mathrm{d}}(x_i, x_k), \hat{\mathrm{d}}(x_j, x_k) \big\} < \tau \big\}$ for some threshold $\tau > 0$. For $x_i \notin \mathcal{C}(x_{\text{new}})$, the distance is estimated as

$$\hat{\mathrm{d}}(x_i, x_{\text{new}}) = \begin{cases} \sum_{x_k \in \mathcal{C}(x_{\text{new}})} \frac{\hat{\mathrm{d}}(x_k, x_i) - \hat{\mathrm{d}}(x_k, x_{\text{new}})}{|\mathcal{C}(x_{\text{new}})|}. & \text{if } x_i \in \mathcal{V}_{\text{obs}} \\ \sum_{(x_k, x_j) \in \mathcal{C}(x_{\text{new}}) \times \mathcal{C}(i)} \frac{\hat{\mathrm{d}}(x_k, x_j) - \hat{\mathrm{d}}(x_k, x_{\text{new}}) - \hat{\mathrm{d}}(x_j, y_i)}{|\mathcal{C}(x_{\text{new}})||\mathcal{C}(i)|} & \text{otherwise} \end{cases}. \tag{9}$$

The set $\mathcal{K}_{ij}$ is designed to ensure that the nodes involved in the calculation of information distances are not too far, since estimating long distances requires a large number of samples. The maximal cardinality of $\mathcal{K}_{ij}$ over all nodes $x_i, x_j \in \mathcal{V}$ can be found, and we denote this as N_τ, i.e., $|\mathcal{K}_{ij}| \leq N_\tau$.

The observed nodes are placed in the 0th layer. The hidden nodes introduced in i^{th} iteration are placed in i^{th} layer. The nodes in the ith layer are in the active set Γ^{i+1} in the $(i+1)^{\text{st}}$ iteration, but nodes in Γ^{i+1} can be nodes created in the jth iteration, where $j < i$. For example, in Fig. 1, nodes x_{12}, x_{14} and x_{15} are created in the 1st iteration, and they are in Γ^2. Nodes x_1, x_2 and x_5 are also in Γ^2, which are observed nodes. Eqns. (8) and (9) imply that the estimation error in the 0th layer will propagate to the nodes in higher layers, and it is necessary to derive concentration results for the information distance related to the nodes in higher layers. To avoid repeating complicated expressions in the various concentration bounds to follow, we define the function

Figure 1: An illustration of the active set. The shaded nodes are the observed nodes and the rest are hidden nodes.

$$f(x) \triangleq 2l_{\max}^2 e^{-\frac{3n_2}{32\lambda\kappa n_1}x} + l_{\max}^2 e^{-c\frac{n_2}{4\lambda^2\kappa^2}x^2} =: ae^{-wx} + be^{-ux^2},$$

where $\lambda = 2l_{\max}^2 e^{\rho_{\max}/l_{\max}} / \delta_{\min}^{1/l_{\max}}$, $w = \frac{3n_2}{32\lambda\kappa n_1}$, $u = c\frac{n_2}{4\lambda^2\kappa^2}$, $a = 2l_{\max}^2$ and $b = l_{\max}^2$. To assess the proximity of the estimates $\hat{\mathrm{d}}(x_i, x_{\text{new}})$ in (8) and (9) to their nominal versions, we define

$$h^{(l)}(x) \triangleq s^l f(m^l x) = s^l \big(ae^{-wm^l x} + be^{-um^{2l}x^2} \big) \quad \text{for all} \quad l \in \mathbb{N} \cup \{0\}. \tag{10}$$

where $s = d_{\max}^2 + 2d_{\max}^3(1 + 2N_\tau)$ and $m = 2/9$. The following proposition yields *recursive estimates* for the errors of the distances at various layers of the learned latent tree.

Proposition 4. *With Assumptions 1–5, if we implement the truncated inner product to estimate the information distance among observed nodes and adopt (8) and (9) to estimate the information distances related to newly introduced hidden nodes, then the information distance related to the hidden nodes x_{new} created in the lth layer $\hat{\mathrm{d}}(x_i, x_{\text{new}})$ satisfies*

$$\mathbb{P}\Big(\big| \hat{\mathrm{d}}(x_i, x_{\text{new}}) - \mathrm{d}(x_i, x_{\text{new}}) \big| > \varepsilon \Big) < h^{(l)}(\varepsilon) \quad \text{for all} \quad x_i \in \Gamma^{l+1} \quad \text{and} \quad l \in \mathbb{N} \cup \{0\}. \tag{11}$$

We note that Proposition 4 demonstrates that the coefficient of exponential terms in (11) grow exponentially with increasing layers (i.e., m^l and m^{2l} in (10)), which requires a commensurately large number of samples to control the tail probabilities.

Theorem 1. *Under Assumptions 1–5, RRG learns the correct latent tree with probability $1 - \eta$ if*

$$n_2 = \Omega\Big(\frac{l_{\max}^4 e^{2\rho_{\max}/l_{\max}} \kappa^2}{\delta_{\min}^{2/l_{\max}} \rho_{\min}^2} \Big(\frac{9}{2}\Big)^{2L_{\text{R}}} \log \frac{|\mathcal{V}_{\text{obs}}|^3}{\eta} \Big) \quad \text{and} \quad n_1 = O\Big(\frac{\sqrt{n_2}}{\log n_2} \Big), \tag{12}$$

where L_{R} is the number of iterations of RRG needed to construct the tree.

Theorem 1 indicates that the number of clean samples n_2 required by RRG to learn the correct structure grows exponentially with the number of iterations L_R. Specifically, for the full m-tree illustrated in Fig. 5, n_2 is exponential in the depth of the tree with high probability for structure learning to succeed. The sample complexity of RRG depends on $e^{2\rho_{\max}/l_{\max}}$, and the exponential relationship with $\rho_{\max}$ will be shown to be unavoidable in view of our impossibility result in Theorem 5. Huang et al. [17, Lemma 7.2] also derived a sample complexity result for learning latent trees but the algorithm is based on [5] instead of RG. RRG is able to tolerate $n_1 = O(\sqrt{n_2}/\log n_2)$ corruptions. This tolerance level originates from the properties of the truncated inner product; similar tolerances will also be seen for the sample complexities of subsequent algorithms. We expect this is also the case for [17], which is based on [5], though we have not shown this formally. In addition, the sample complexity is applicable to a wide class of graphical models that satisfies the Assumptions 1 to 5, while the sample complexity result [4, Theorem 11], which hides the dependencies on the parameters, only holds for a limited class of graphical models whose effective depths (the maximal length of paths between hidden nodes and their closest observed nodes) are bounded in $|\mathcal{V}_{\text{obs}}|$.

3.3 Robust Neighbor Joining and Spectral Neighbor Joining algorithms

The NJ algorithm [1] also makes use of additive distances to identify the existence of hidden nodes. To robustify the NJ algorithm, we adopt robust estimates of information distances as the additive distances in the so-called RNJ algorithm. We first recap a result by Atteson [18].

Proposition 5. *If all the nodes have exactly two children, NJ will output the correct latent tree if*

$$\max_{x_i,x_j\in\mathcal{V}_{\text{obs}}}\left|\hat{d}(x_i,x_j)-d(x_i,x_j)\right|\leq\rho_{\min}/2. \tag{13}$$

Unlike RG, NJ does not identify the parent relationship among nodes, so it is only applicable to binary trees in which each node has at most two children.

Theorem 2. *If Assumptions 1–5 hold and all the nodes have exactly two children, RNJ constructs the correct latent tree with probability at least $1-\eta$ if*

$$n_2=\Omega\Big(\frac{l_{\max}^4 e^{2\rho_{\max}/l_{\max}}\kappa^2}{\delta_{\min}^{2/l_{\max}}\rho_{\min}^2}\log\frac{|\mathcal{V}_{\text{obs}}|^2}{\eta}\Big)\quad and\quad n_1=O\Big(\frac{\sqrt{n_2}}{\log n_2}\Big). \tag{14}$$

Theorem 2 indicates that the sample complexity of RNJ grows as $\log|\mathcal{V}_{\text{obs}}|$, which is much better than RRG. Similarly to RRG, the sample complexity has an exponential dependence on $\rho_{\max}$.

In recent years, several variants of NJ algorithms have been proposed. The additivity of information distances results in certain properties of the rank of the matrix $\mathbf{R}\in\mathbb{R}^{|\mathcal{V}_{\text{obs}}|\times|\mathcal{V}_{\text{obs}}|}$, where $\mathbf{R}(i,j)=\exp(-d(x_i,x_j))$ for all $x_i,x_j\in\mathcal{V}_{\text{obs}}$. Jaffe *et al.* [6] proposed SNJ which utilizes the rank of $\mathbf{R}$ to deduce the sibling relationships among nodes. We robustify the SNJ algorithm by implementing the robust estimation of information distances, as shown in Algorithm 2.

Although SNJ was designed for discrete random variables, the additivity of the information distance proved in Proposition 1 guarantees the consistency of Robust Spectral NJ (RSNJ) for GGMs with vector variables. A sufficient condition for RSNJ to learn the correct tree can be generalized from [6].

Proposition 6. *If Assumptions 1–5 hold and all the nodes have exactly two children, a sufficient condition for RSNJ to recover the correct tree from $\hat{\mathbf{R}}$ is*

$$\|\hat{\mathbf{R}}-\mathbf{R}\|_2\leq g(|\mathcal{V}_{\text{obs}}|,\rho_{\min},\rho_{\max}), \tag{15}$$

where

$$g(x,\rho_{\min},\rho_{\max})=\begin{cases}\frac{1}{2}(2e^{-\rho_{\max}})^{\log_2(x/2)}e^{-\rho_{\max}}\big(1-e^{-2\rho_{\min}}\big), & e^{-2\rho_{\max}}\leq 0.5\\ e^{-3\rho_{\max}}\big(1-e^{-2\rho_{\min}}\big), & e^{-2\rho_{\max}}>0.5\end{cases}.$$

Similar with RNJ, RSNJ also does not identify the parent relationship between nodes, so it only applies to binary trees. To state the next result succinctly, we assume that $\rho_{\max}\geq\frac{1}{2}\log 2$; this is the regime of interest because we consider large trees which implies that $\rho_{\max}$ is typically large.

Theorem 3. *If Assumptions 1–5 hold, $\rho_{\max}\geq\frac{1}{2}\log 2$, and all the nodes have exactly two children, RSNJ reconstructs the correct latent tree with probability at least $1-\eta$ if*

$$n_2=\Omega\Big(\frac{l_{\max}^4 e^{2\rho_{\max}(1/l_{\max}+\log_2(|\mathcal{V}_{\text{obs}}|/2)+1)}\kappa^2}{\delta_{\min}^{2/l_{\max}}e^{2\rho_{\min}}}\log\frac{|\mathcal{V}_{\text{obs}}|^2}{\eta}\Big)\quad and\quad n_1=O\Big(\frac{\sqrt{n_2}}{\log n_2}\Big). \tag{16}$$

Theorem 3 indicates that the sample complexity of RSNJ grows as $\text{poly}(|\mathcal{V}_{\text{obs}}|)$. Specifically, in the binary tree case, the sample complexity grows exponentially with the depth of the tree. Also, the dependence of sample complexity on $\rho_{\max}$ is exponential, i.e., $O\big(e^{2(1/l_{\max}+\log_2(|\mathcal{V}_{\text{obs}}|/2)+1)\rho_{\max}}\big)$, but the coefficient of $\rho_{\max}$ is larger than those of RRG and RNJ, which are $O\big(e^{2\rho_{\max}/l_{\max}}\big)$. Compared to the sample complexity of SNJ in [6], the sample complexity of RSNJ has the same dependence on the number of observed nodes $|\mathcal{V}_{\text{obs}}|$, which means that the robustification of SNJ using the truncated inner product is able to tolerate $O\big(\frac{\sqrt{n_2}}{\log n_2}\big)$ corruptions.

3.4 Robust Chow-Liu Recursive Grouping

In this section, we show that the exponential dependence on L_{R} in Theorem 1 can be provably mitigated with an accurate initialization of the structure. Different from RRG, RCLRG takes Chow-Liu algorithm as the initialization stage, as shown in Algorithm 3. The Chow-Liu algorithm [19] learns the maximum likelihood estimate of the tree structure by finding the maximum weight spanning tree of the graph whose edge weights are the mutual information quantities between these variables. With the monotonicity between mutual information and information distance, we can construct the Chow-Liu tree as the minimum spanning tree (MST) of the graph whose weights are information distances. However, this monotonicity property is violated in general for random *vectors*, since the determinants of $\mathbf{I} - \mathbf{X}$ and $\mathbf{X}$ are not monotonic functions of each other for general $\mathbf{X} \in \mathbb{R}^{l_{\max} \times l_{\max}}$. We now show that the required monotonicity property is guaranteed by a wide class of GGMs.

We choose any node in the tree as the root node $\mathbf{x}_{\text{r}}$, and define the parent node and set of children nodes (in the rooted tree) of any node x_i as $\text{pa}(i)$ and $\mathcal{C}(x_i)$ respectively. The depth of a node x_i is $\text{d}_{\mathbb{T}}(x_i, x_{\text{r}})$. We specify the model where

$$\mathbf{x}_i = \mathbf{A}\mathbf{x}_{\text{pa}(i)} + \mathbf{n}_i \quad \text{for all} \quad x_i \in \mathcal{V} \tag{17}$$

where $\mathbf{A} \in \mathbb{R}^{l_{\max} \times l_{\max}}$ is non-singular, $\mathbf{n}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{\Sigma}_i)$ and $\mathbf{n}_i$'s are mutually independent. Since the root node has no parent, it is natural to set $\mathbf{x}_{\text{pa}(\text{r})} = \mathbf{0}$ and $\mathbf{n}_{\text{r}} \sim \mathcal{N}(\mathbf{0}, \mathbf{\Sigma}_{\text{r}})$. It is easy to verify that the model specified by (17) and this initial condition is an undirected GGM. To guarantee that the mutual information is a monotonic function of the information distance, it is natural to consider the situation in which the covariance matrices of all variables are the same up to a constant scale factor.

Proposition 7. *If $\mathbf{n}_i$'s for the variables at depth l are distributed as $\mathcal{N}(0, \alpha^{l-1}\mathbf{\Sigma}_{\text{n}})$, and*

$$\mathbf{A}\mathbf{\Sigma}_{\text{r}}\mathbf{A}^{\text{T}} + \mathbf{\Sigma}_{\text{n}} = \alpha\mathbf{\Sigma}_{\text{r}} \tag{18}$$

where $\alpha > 0$ is a constant, then the covariance matrix of the variable at depth l is $\alpha^l\mathbf{\Sigma}_{\text{r}}$.

We call (18) the $(\mathbf{A}, \mathbf{\Sigma}_{\text{r}}, \mathbf{\Sigma}_{\text{n}})$-*homogenous* condition, which guarantees that covariance matrices of the random vectors in the tree are same up to a scale factor. We now provide a sufficient condition on $\mathbf{A}, \mathbf{\Sigma}_{\text{r}}, \mathbf{\Sigma}_{\text{n}}$ to achieve the monotonicity between mutual information and information distance.

Proposition 8. *In the GGM specified by (17), if (i) the $\mathbf{n}_i$'s for the variables at depth l are distributed as $\mathcal{N}(\mathbf{0}, \alpha^{l-1}\mathbf{\Sigma}_{\text{n}})$ for some $\alpha > 0$; (ii) the $(\mathbf{A}, \mathbf{\Sigma}_{\text{r}}, \mathbf{\Sigma}_{\text{n}})$-homogeneous condition in (18) is satisfied; and (iii) $\mathbf{\Sigma}_{\text{r}}$ and $\mathbf{A}$ commute, then the mutual information is a monotonically decreasing function of the information distance. Furthermore, the mutual information and the information distance can be expressed in closed form in terms of $(\alpha, \text{d}_{\mathbb{T}}(x_i, x_j), \mathbf{A})$. See (J.27) and (J.28) in Appendix J.*

This property is trivially satisfied in the scalar case [4], but is more subtle in the vector case. With this property, the Chow-Liu algorithm [19] can be implemented by finding the MST with information distances as edge weights.

Lemma 9. *If the three conditions in Proposition 8 are satisfied, the Chow-Liu tree reduces to the MST where edge weights are the information distances, i.e.,*

$$\mathbb{T}_{\text{CL}} = \text{MST}(\mathcal{V}_{\text{obs}}; \mathbf{D}) := \underset{\mathbb{T} \in \mathcal{T}_{\mathcal{V}_{\text{obs}}}}{\arg\min} \sum_{(x_i, x_j) \in \mathbb{T}} \text{d}(x_i, x_j), \tag{19}$$

where $\mathcal{T}_{\mathcal{V}_{\text{obs}}}$ is the set of all the trees with node set $\mathcal{V} = \mathcal{V}_{\text{obs}}$.

Definition 2. *Given the latent tree $\mathbb{T} = (\mathcal{V}, \mathcal{E})$ and any node $x_i \in \mathcal{V}$, the surrogate node [4] of x_i is* $\text{Sg}(x_i; \mathbb{T}, \mathcal{V}_{\text{obs}}) = \arg\min_{x_j \in \mathcal{V}_{\text{obs}}} \text{d}(x_i, x_j).$

We introduce a new notion of distance that quantifies the sample complexity of RCLRG.

Definition 3. *Given the latent tree* $\mathbb{T} = (\mathcal{V}, \mathcal{E})$ *and any node* $x_i \in \mathcal{V}$*, the* contrastive distance *of* x_i *with respect to* $\mathcal{V}_{\mathrm{obs}}$ *is defined as*

$$\mathrm{d}_{\mathrm{ct}}(x_i; \mathbb{T}, \mathcal{V}_{\mathrm{obs}}) = \min_{x_j \in \mathcal{V}_{\mathrm{obs}} \setminus \mathrm{Sg}(x_i; \mathbb{T}, \mathcal{V}_{\mathrm{obs}})} \mathrm{d}(x_i, x_j) - \min_{x_j \in \mathcal{V}_{\mathrm{obs}}} \mathrm{d}(x_i, x_j). \tag{20}$$

Definitions 2 and 3 imply that the surrogate node $\mathrm{Sg}(x_i; \mathbb{T}, \mathcal{V}_{\mathrm{obs}})$ of any observed node x_i is itself x_i, and its contrastive distance is the information distance between the closest observed node and itself. It is shown that the Chow-Liu tree $\mathbb{T}_{\mathrm{CL}}$ is equal to the tree where all the hidden nodes are contracted to their surrogate nodes [4], so it will be difficult to identify the surrogate node of some node if its contrastive distance is small. Under this scenario, more accurate estimates of the information distances are required to construct the correct Chow-Liu tree.

Proposition 10. *The Chow-Liu tree* $\mathrm{MST}(\mathcal{V}_{\mathrm{obs}}; \hat{\mathbf{D}})$ *is constructed correctly if*

$$\left| \hat{\mathrm{d}}(x_i, x_j) - \mathrm{d}(x_i, x_j) \right| < \Delta_{\mathrm{MST}}/2 \quad \textit{for all} \quad x_i, x_j \in \mathcal{V}_{\mathrm{obs}}, \tag{21}$$

where $\Delta_{\mathrm{MST}} := \min_{x_j \in \mathrm{Int}(\mathbb{T})} \mathrm{d}_{\mathrm{ct}}(x_j; \mathbb{T}, \mathcal{V}_{\mathrm{obs}})$.

Hence, the contrastive distance describes the difficulty of learning the correct Chow-Liu tree.

Theorem 4. *With Assumptions 1–4 and the conditions of Proposition 8, RCLRG constructs the correct latent tree with probability at least* $1 - \eta$ *if*

$$n_2 = \Omega\left(\max\left\{ \frac{1}{\rho_{\min}^2} \left(\frac{9}{2}\right)^{2L_{\mathrm{C}}}, \frac{1}{\Delta_{\mathrm{MST}}^2} \right\} \frac{l_{\max}^4 e^{2\rho_{\max}/l_{\max}} \kappa^2}{\delta_{\min}^{2/l_{\max}}} \log \frac{|\mathcal{V}_{\mathrm{obs}}|^3}{\eta} \right) \quad \textit{and} \quad n_1 = O\left(\frac{\sqrt{n_2}}{\log n_2} \right), \tag{22}$$

where L_{C} *is the maximum number of iterations of RRG (over each internal node of the constructed Chow-Liu tree) in RCLRG needed to construct the tree.*

If we implement RCLRG with *true* information distances, $L_{\mathrm{C}} \leq \lceil \frac{1}{2}\mathrm{Deg}(\mathrm{MST}(\mathcal{V}_{\mathrm{obs}}; \hat{\mathbf{D}})) - 1 \rceil$. Theorem 4 indicates that the sample complexity of RCLRG grows exponentially in $L_{\mathrm{C}} \ll L_{\mathrm{R}}$. Compared with [4, Theorem 12], the sample complexity of RCLRG in Theorem 4 is applicable to a wide class of graphical models that satisfy Assumptions 1 to 5, while the [4, Theorem 12] requires the assumption that the effective depths of latent trees are *bounded* in $|\mathcal{V}_{\mathrm{obs}}|$, which is rather restrictive.

3.5 Comparison of robust latent tree learning algorithms

Since the sample complexities of RRG, RCLRG, RSNJ and RNJ depend on different parameters and different structures of the underlying graphs, it is instructive to compare the sample complexities of these algorithms on some representative tree structures. These trees are illustrated in Fig. 5. RSNJ and RNJ are not able to identify the parent relationship among nodes, so they are only applicable to trees whose maximal degrees are no larger that 3, including the double-binary tree and the HMM. In particular, RNJ and RSNJ are not applicable to the full m-tree (for $m \geq 3$) and the double star. Derivations and more detailed discussions of the sample complexities are deferred to Appendix K.

| n_2 Algorithm
Tree	RRG	RCLRG	RSNJ	RNJ
Double-binary tree	$O\big(\psi(\frac{9}{2})^{\mathrm{Diam}(\mathbb{T})}\big)$	$O\big(\psi(\frac{9}{2})^{\frac{1}{2}\mathrm{Diam}(\mathbb{T})}\big)$	$O\big(e^{2t\rho_{\max}}\mathrm{Diam}(\mathbb{T})\big)$	$O\big(\psi\mathrm{Diam}(\mathbb{T})\big)$
HMM	$O\big(\psi(\frac{9}{2})^{\mathrm{Diam}(\mathbb{T})}\big)$	$O\big(\psi \log \mathrm{Diam}(\mathbb{T})\big)$	$O\big(e^{2t\rho_{\max}} \log \mathrm{Diam}(\mathbb{T})\big)$	$O\big(\psi \log \mathrm{Diam}(\mathbb{T})\big)$
Full m-tree	$O\big(\psi(\frac{9}{2})^{\mathrm{Diam}(\mathbb{T})}\big)$	$O\big(\psi\mathrm{Diam}(\mathbb{T})\big)$	N.A.	N.A.
Double star	$O\big(\psi \log d_{\max}\big)$	$O\big(\psi \log d_{\max}\big)$	N.A.	N.A.

Table 1: The sample complexities of RRG, RCLRG, RSNJ and RNJ on the double-binary tree, the HMM, the full m-tree and the double star. We set $\psi := e^{2\rho_{\max}/l_{\max}}$ and $t = O(l_{\max}^{-1} + \log |\mathcal{V}_{\mathrm{obs}}|)$.

3.6 Experimental results

We present simulation results to demonstrate the efficacy of the robustified algorithms. Samples are generated from a HMM with $l_{\max} = 3$ and $\mathrm{Diam}(\mathbb{T}) = 80$. The three conditions in Proposition 8 are satisfied with $\alpha = 1$. The Robinson-Foulds distance [20] between the true and estimated trees is

320 adopted to measure the performances of the algorithms. For the implementations of CLRG and RG,
321 we use the code from [4]. Other settings and more extensive experiments are given in Appendix L.

322 We consider three corruption patterns here. (i) *Uniform corruptions* are independent additive noises
323 in $[-2A, 2A]$; (ii) *Constant magnitude corruptions* are also independent additive noises but taking
324 values in $\{-A, +A\}$ with probability 0.5. These two types of noises are distributed randomly in $\mathbf{X}_1^n$;
325 (iii) *HMM corruptions* are generated by a HMM which has the same structure as the original HMM
326 but has different parameters. They replace the entries in $\mathbf{X}_1^n$ with samples generated by the variables
327 in the same positions. In our simulations, A is set to 60, and the number of corruptions n_1 is 100.

(a) Uniform corruptions   (b) Constant magnitude corruptions   (c) HMM corruptions

Figure 2: Robinson-Foulds distances of robustified and original algorithms averaged over 100 trials

328 Fig. 2 (error bars are in Appendix L.1) demonstrates the superiority of RCLRG in learning HMMs
329 compared to other algorithms. The robustified algorithms also result in smaller estimation errors
330 (Robinson-Foulds distances) compared to their unrobustified counterparts in presence of corruptions.

4   Impossibility result

Definition 4. *Given a triple* $(|\mathcal{V}_{\mathrm{obs}}|, \rho_{\max}, l_{\max})$, *the set* $\mathcal{T}(|\mathcal{V}_{\mathrm{obs}}|, \rho_{\max}, l_{\max})$ *consists of all multi-
variate Gaussian distributions* $\mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$ *such that: (1) The underlying graph* $\mathbb{T} = (\mathcal{V}, \mathcal{E})$ *is a tree*
$\mathbb{T} \in \mathcal{T}_{\geq 3}$, *and the size of the set of observed nodes is* $|\mathcal{V}_{\mathrm{obs}}|$. *(2) The distribution* $\mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$ *satisfies
Assumptions 1 and 5 with parameters* $l_{\max}$ *and* $\rho_{\max}$.

336 For the given class of graphical models $\mathcal{T}(|\mathcal{V}_{\mathrm{obs}}|, \rho_{\max}, l_{\max})$, nature chooses some parameter $\theta = \boldsymbol{\Sigma}$
337 and generates n i.i.d. samples $\mathbf{X}_1^n$ from $\mathbb{P}_\theta$. The goal of the statistician is to use the observations $\mathbf{X}_1^n$
338 to learn the underlying graph $\mathbb{T}$, which entails the design of a decoder $\phi : \mathbb{R}^{n|\mathcal{V}_{\mathrm{obs}}|l_{\max}} \to \mathcal{T}_{\geq|\mathcal{V}_{\mathrm{obs}}|}$,
339 where $\mathcal{T}_{\geq|\mathcal{V}_{\mathrm{obs}}|}$ is the set of trees whose size of the node set is at least $|\mathcal{V}_{\mathrm{obs}}|$.

Theorem 5. *Consider the class of graphical models* $\mathcal{T}(|\mathcal{V}_{\mathrm{obs}}|, \rho_{\max}, l_{\max})$, *where* $|\mathcal{V}_{\mathrm{obs}}| \geq 3$. *If
there exists a graph decoder learns from* n *i.i.d. samples such that*

$$\max_{\theta(\mathbb{T}) \in \mathcal{T}(|\mathcal{V}_{\mathrm{obs}}|, \rho_{\max}, l_{\max})} \mathbb{P}_{\theta(\mathbb{T})}(\phi(\mathbf{X}_1^n) \neq \mathbb{T}) < \delta, \tag{23}$$

then (as $\rho_{\max} \to \infty$ *and* $|\mathcal{V}_{\mathrm{obs}}| \to \infty$),

$$n = \max \left\{ \Omega\left((1-\delta) e^{\frac{\rho_{\max}}{\lfloor \log_3 |\mathcal{V}_{\mathrm{obs}}| \rfloor l_{\max}}} \log |\mathcal{V}_{\mathrm{obs}}|\right), \Omega\left((1-\delta) e^{\frac{2\rho_{\max}}{3 l_{\max}}}\right) \right\}. \tag{24}$$

343 Theorem 5 implies that the optimal sample complexity grows as $\Omega(\log |\mathcal{V}_{\mathrm{obs}}|)$ as $|\mathcal{V}_{\mathrm{obs}}|$ grows. Table 1
344 indicates that the sample complexity of RCLRG when the underlying latent tree is a full m-tree (for
345 $m \geq 3$) or a HMM is optimal in the dependence on $|\mathcal{V}_{\mathrm{obs}}|$. The sample complexity of RNJ is also
346 optimal in $|\mathcal{V}_{\mathrm{obs}}|$ for double binary trees and HMMs. In contrast, the derived sample complexities of
347 RRG and RSNJ are suboptimal in relation to Theorem 5. However, one caveat of our analyses of the
348 latent tree learning algorithms in Section 3 is that we are not claiming that they are the best possible
349 for the given algorithm; there may be room for improvement.

350 When the maximum information distance $\rho_{\max}$ grows, Theorem 5 indicates that the optimal sample
351 complexity grows as $\Omega(e^{\frac{2\rho_{\max}}{3 l_{\max}}})$. Table 1 shows the sample complexities of RRG, RCLRG and RNJ
352 grow as $O(e^{2\frac{\rho_{\max}}{l_{\max}}})$, which has the alike dependence as the impossibility result. However, the sample
353 complexity of RSNJ grows as $O(e^{2t\rho_{\max}})$, which is larger (looser) than that prescribed by Theorem 5.

References

[1] N. Saitou and M. Nei. The neighbor-joining method: a new method for reconstructing phyloge-
netic trees. *Mol. Bio. Evol.*, 4(4):406–425, 1987.

[2] D. Tang, H. J. Chang, A. Tejani, and T. Kim. Latent regression forest: Structured estimation of
3d articulated hand posture. In *Proc. IEEE Conf. Computer Vision and Pattern Recognition*,
pages 3786–3793, 2014.

[3] J. Eisenstein, B. O'Connor, N. A. Smith, and E. Xing. A latent variable model for geographic
lexical variation. In *Proc. Conf. Empirical Methods in Natural Language Processing*, pages
1277–1287, 2010.

[4] M. J. Choi, V. Y. F. Tan, A. Anandkumar, and A. S. Willsky. Learning latent tree graphical
models. *Journal of Machine Learning Research*, 12:1771–1812, 2011.

[5] A. Anandkumar, K. Chaudhuri, D. Hsu, S. M. Kakade M, L. Song, and T. Zhang. Spectral
methods for learning multivariate latent tree structure. *arXiv preprint arXiv:1107.1283*, 2011.

[6] A. Jaffe, N. Amsel, Y. Aizenbud, B. Nadler, J. T. Chang, and Y. Kluger. Spectral neighbor
joining for reconstruction of latent tree models. *SIAM Journal on Mathematics of Data Science*,
3(1):113–141, 2021.

[7] M. Casanellas, M. Garrote-Lopez, and P. Zwiernik. Robust estimation of tree structured models.
arXiv:2102.05472v1 [stat.ML], Feb. 2021.

[8] A. Katiyar, J. Hoffmann, and C. Caramanis. Robust estimation of tree structured gaussian
graphical models. In *International Conference on Machine Learning*, pages 3292–3300. PMLR,
2019.

[9] K. E. Nikolakakis, D. S. Kalogerias, and A.D. Sarwate. Learning tree structures from noisy
data. In *Proc. Artificial Intelligence and Statistics*, pages 1771–1782. PMLR, 2019.

[10] A. Tandon, V. Y. F. Tan, and S. Zhu. Exact asymptotics for learning tree-structured graphical
models: Noiseless and noisy samples. *IEEE Journal on Selected Areas of Information Theory*,
1(3):760–776, 2020.

[11] A. Tandon, A. H. J. Yuan, and V. Y. F. Tan. SGA: A robust algorithm for partial recovery of
tree-structured graphical models with noisy samples. In *International Conference on Machine
Learning*. PMLR, 2021.

[12] L. Wang and Q. Gu. Robust gaussian graphical model estimation with arbitrary corruption. In
International Conference on Machine Learning, pages 3617–3626. PMLR, 2017.

[13] A. P. Parikh, L. Song, and E. P. Xing. A spectral algorithm for latent tree graphical models. In
International Conference on Machine Learning, pages 1065–1072, Jun. 2011.

[14] Y. Chen, C. Caramanis, and S. Mannor. Robust high dimensional sparse regression and matching
pursuit. *arXiv preprint arXiv:1301.2725*, 2013.

[15] S. L. Lauritzen. *Graphical models*, volume 17. Clarendon Press, 1996.

[16] P. Judea. *Probabilistic reasoning in intelligent systems: Networks of plausible inference*.
Elsevier, 2014.

[17] F. Huang, N. U. Naresh, I. Perros, R. Chen, J. Sun, and A. Anandkumar. Guaranteed scalable
learning of latent tree models. In *Uncertainty in Artificial Intelligence*, pages 883–893. PMLR,
2020.

[18] K. Atteson. The performance of neighbor-joining methods of phylogenetic reconstruction.
Algorithmica, 25(2):251–278, 1999.

[19] C. Chow and C. Liu. Approximating discrete probability distributions with dependence trees.
IEEE Trans. Inform. Theory, 14(3):462–467, 1968.

[20] D. F. Robinson and L. R. Foulds. Comparison of phylogenetic trees. *Mathematical Biosciences*, 53(1-2):131–147, 1981.

[21] R. Vershynin. *Introduction to the non-asymptotic analysis of random matrices*. Cambridge University Press, New York, NY, 2010.

[22] G. H. Golub and C. F. Van Loan. *Matrix Computations*. The Johns Hopkins University Press, Maryland, US, 2013.

[23] G. W. Stewart. Perturbation theory for the singular value decomposition. *SVD and Signal Processing, II: Algorithms, Analysis and Applications*, pages 99–109, 1991.

[24] R. A. Horn and C. R. Johnson. *Matrix analysis*. Cambridge University Press, New York, NY, 2012.

[25] S. Pettie and V. Ramachandran. An optimal minimum spanning tree algorithm. *J. ACM*, 49(1):16–34, 2002.

[26] W. Wang, M. J. Wainwright, and K. Ramchandran. Information-theoretic bounds on model selection for Gaussian markov random fields. In *Proc. IEEE Int. Symp. on Inf. Theory*, pages 1373–1377, Austin, Texas, USA, Jun. 2010. IEEE.

[27] M. Marcus and W. Gordon. An extension of the Minkowski determinant theorem. *Proceedings of the Edinburgh Mathematical Society*, 17(4):321–324, 1971.